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1 **Local trade-offs shape flower size evolution** 2 **across *Arabidopsis thaliana* distribution**

3
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16 **Author Contributions:**

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18 KS and CFM; Materials and analysis tools contributions: DV and JRL; writing original draft KS and AS;
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22 The authors declare no competing interests.

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25

26

27 **Abstract**

28 The diversification of flowers, largely driven by mutualistic interactions with animal pollinators, has
29 generated remarkable variation in floral form and size and is thought to have driven the evolutionary
30 radiation of angiosperms. Here, we investigate the geographic variation in reproductive organ growth
31 in the self-fertilizing species *Arabidopsis thaliana*, where the loss of pollinator dependence is predicted
32 to favour reduced floral investment through resource reallocation. We find extensive variation in flower
33 size underpinned by a polygenic architecture, in which derived alleles both increase and decrease petal
34 size and show signatures of positive selection. The direction of flower-size evolution varies
35 geographically, reflecting environmentally mediated shifts in the relationship between flower size and
36 seed production. Strong purifying selection at the species' climatic margins favors smaller flowers,
37 whereas relaxed environmental constraints in more suitable habitats allow the persistence and potential
38 adaptive value of large-flower variants. This biogeographic pattern extends to other life-history traits,
39 illustrating how environmental heterogeneity, acting through resource allocation constraints, shapes
40 evolutionary trajectories and maintains phenotypic diversity in selfing lineages.

41

42

43 Introduction

44 Natural selection balances flower organ construction and maintenance costs against the likelihood of
45 fertilization (Roddy et al., 2020). These costs follow classical allocation theory, as illustrated by the
46 trade-off between flower size and number (Roddy et al., 2020)(Caruso, 2006; Worley & Barrett, 2000).
47 In animal-pollinated outcrossing species, pollinator-mediated selection often favors large and more
48 conspicuous flowers adapted to pollinator nutritional needs and behaviours, promoting flower
49 diversification (Crane et al., 1994; Friis et al., 1994, 1999, 2006; Galen, 1999). Following the transitions
50 to self-fertilization, however, selection is expected to reduce investment in floral displays and reallocate
51 resources to other fitness components (Stebbins, 1957). This mechanism is frequently invoked to
52 explain the characteristic reduction in flower size and attractiveness, collectively known as the “selfing
53 syndrome”, in selfing species (Sicard et al., 2011). Yet, empirical studies quantifying the relative
54 importance of resource-allocation trade-offs compared to other evolutionary processes remain limited.

55 Different evolutionary processes could influence floral evolution in selfing species, each leading to
56 distinct patterns of variation in flower size. When the construction of floral organs incurs significant
57 energetic costs, selection should favor mutations that reduce floral display size, thereby enhancing seed
58 production through resource reallocation (Charlesworth & Charlesworth, 1981). This directional
59 selection scenario predicts (i) a negative association between flower size and seed production, (ii)
60 genomic evidence of positive selection on alleles that reduce flower size, and (iii) a general convergence
61 toward smaller flowers with reduced within-species variance (Stephan, 2019). Because environmental
62 heterogeneity influences both resource availability and selection strength (Troth et al., 2018), these
63 patterns should be more pronounced in resource-limited environments (Fu et al., 2010; MacTavish &
64 Anderson, 2020; Tuller et al., 2018). Alternatively, the loss of attractive traits could confers little to no
65 fitness benefit and selection maintaining large displays would be relaxed. Under this scenario the
66 trajectory of floral-size evolution may become less predictable. In this case, the role of genetic drift
67 should play a more dominant role, leading to idiosyncratic shifts in mean size, widening phenotypic
68 variance, and few clear signatures of directional selection at causal loci (Woźniak & Sicard, 2018).
69 Reduced effective recombination under high homozygosity can further facilitate drift-mediated fixation
70 of nonsynonymous mutations that impair floral-growth genes (Barrett et al., 2014).

71 Additional processes may interact with these trajectories. Pleiotropy can couple floral size variation with
72 other organs controlled by shared developmental pathways, producing correlated responses that range
73 from near-isometric to clearly allometric depending on the extent of shared genetic control (Huxley et
74 al., 1993; Johnson & Lenhard, 2011) . Likewise, linked selection, enhanced in selfers due to low
75 effective recombination, can further reshape diversity around selected loci and indirectly hitchhike
76 alleles affecting floral size (Slotte, 2014). Given the diversity and nature of the mechanisms involved,
77 floral-size variation is likely to be highly polygenic, consistent with observations in selfing lineages
78 (Abraham et al., 2013; Sicard et al., 2011; Sicard & Lenhard, 2011; Woźniak et al., 2020). However, the
79 relative contributions of these mechanisms remain poorly understood.

80 *Arabidopsis thaliana* provides a powerful system to test these predictions. The evolution of self-
81 fertilization in this species occurred relatively recently and was followed by rapid range expansion
82 across diverse climates and environments (Bechsgaard et al., 2006; Shimizu & Tsuchimatsu, 2015;
83 Tang et al., 2007). This history provides opportunities for environment-specific selection on floral
84 investment and the fixation of distinct size-affecting mutations across lineages (Tsuchimatsu et al.,
85 2020). Leveraging extensive genomic and phenotypic resources, we examine the genetic and
86 ecological drivers of floral-size evolution following the emergence of selfing and assess how
87 environmental gradients shape these evolutionary trajectories across the species' range. Our results
88 reveal that patterns of phenotypic evolution differ geographically, reflecting variation in selection
89 pressures across environments.

90

91 **Results**

92 ***Flower size is highly variable within A. thaliana***

93 The directional selection hypothesis predicts that both the mean and variance of flower size decline in
94 self-fertilizing species such as *Arabidopsis thaliana* due to resource allocation trade-offs that favour
95 reduced investment in floral display. Alternatively, genetic and phenotypic variation may persist if drift
96 predominates (Buysse et al., 2024) or if spatially variable environments maintain divergent selection
97 (MacTavish & Anderson, 2020).

98 To test these predictions, we measured the length, width, and area of floral organs and leaves in 407
99 *A. thaliana* accessions spanning the species' global genetic and geographic diversity (**Fig. S1**; **Dataset**
100 **S1**). All traits were variable and approximately normally distributed (**Fig. S2**). Floral organ size explained
101 most phenotypic variance among accessions (43.2%, PC1; **Fig. 1a**), with strong correlations among
102 floral traits ($r > 0.5$). Leaf size variation dominated PC2 (21.6% of the variance). Ovule number
103 contributed mainly to PC4 (<8% of variance). Petal area showed the greatest variability, matching leaf
104 area in its coefficient of variation and spanning more than a threefold range (**Dataset S2**).

105 These results indicate that floral-size variation in *A. thaliana* is substantial despite selfing, consistent
106 with either incomplete fixation of size-reducing alleles during the evolution of selfing or environmentally
107 structured selection maintaining divergence among lineages.

108 ***Independent, low-pleiotropy mutations largely account for the variation in flower organ size.***

109 We investigated whether indirect effects, such as linked selection (Smith & Haigh, 1974) or pleiotropy
110 with other traits under selection, drive the evolution of floral size in *Arabidopsis thaliana*. To this end,
111 we first tested for phenotypic covariation between floral organ size and other traits among genotypes
112 (**Dataset S2**). Stamen, petal, and sepal size were significantly and positively correlated ($r = 0.28$ – 0.93)
113 but showed no correlation with ovule number or leaf size, except for a weak negative correlation
114 between leaf length and petal width ($r = -0.15$; Fig. 1b). The relationship between long-stamen length
115 and petal length was strictly isometric (slope = 1, P-value > 0.01), suggesting an overlap in their genetic

116 bases, while other pairwise relationships were allometric (slope < 1, P-value < 0.001). All floral organs
117 were positively correlated with flowering time ($r = 0.14\text{--}0.42$). We further examined correlations with
118 published phenotypes. Floral organ size traits showed no significant correlation with any of the 107
119 phenotypes from Atwell et al. (2010) (**Dataset S2**). Specific leaf area (SLA), leaf nitrogen content (LNC),
120 and total plant biomass (Przybylska et al. 2023) were significantly correlated with several floral size
121 traits ($r = -0.35\text{--}0.4$), and most traits were also correlated with flowering time ($r = -0.67\text{--}0.70$) in this
122 dataset. To determine whether phenotypic correlations result from shared genetic bases, we quantified
123 genetic correlations using multivariate GWAS (mGWAS) for significantly correlated trait pairs. The
124 mGWAS partitions genetic variance into trait-specific and shared components, allowing estimation of
125 the genetic correlation (r_g) between traits (Zhou & Stephens, 2014) (**Dataset S3**). Within our dataset,
126 we detected no significant genetic correlations between floral and leaf traits, indicating independent
127 genetic bases for their size evolution. Petal and sepal area showed weak but significant correlations
128 with LNC and plant dry mass ($r_g \sim 0.3$, $P < 0.01$), proxies for resource acquisition ability (Pérez-
129 Harguindeguy et al., 2013). These results suggest that resource acquisition capacity modestly
130 influences both vegetative carbon economy and floral organ growth, supporting a role for allocation
131 constraints in flower-size evolution. Overall, floral organ size variation showed limited phenotypic and
132 genetic correlation with other traits, indicating that pleiotropy and linked selection exert only minor
133 influence on floral size evolution in *A. thaliana*.

134 We next conducted univariate GWAS to identify loci underlying individual trait variation (**Fig. 1c**).
135 Despite high SNP-based heritability (PVE = 0.55–0.85) across all floral traits, only three genome-wide
136 significant peaks surpassed the Bonferroni threshold for sepal, petal, and stamen size, consistent with
137 a largely polygenic basis. Two peaks were organ-specific, while one was shared between petal and
138 sepal traits (**Fig. 1c**). To capture additional polygenic signals, we applied a relaxed significance
139 threshold ($P < 5.2 \times 10^{-5}$; see Methods) based on established associations. After filtering for linkage, we
140 identified 61, 105, 87, 59, and 38 SNPs, corresponding to 144, 191, 194, 126, and 117 candidate genes,
141 associated with leaf, petal, sepal, stamen size, and ovule number, respectively (**Fig. 1d**; **Dataset S4**).
142 Functional annotation using TAIR gene ontology terms (Huala et al., 2001) identified several known
143 growth regulators among candidate genes, including *JAGGED*, *GROWTH-REGULATING FACTOR 6*,
144 *CYTOCHROME P450 90B1*, *KINASE-INDUCIBLE DOMAIN INTERACTING 9*, *E2FE*, *MYB DOMAIN*
145 *PROTEIN 104*, *RESPONSE REGULATOR 7*, *INDOLE-3-ACETIC ACID INDUCIBLE 19*, and
146 *SEEDSTICK* (Sauret-Güeto et al., 2013; Omidbakhshfard et al., 2015; Gómez-Ocampo et al., 2023;
147 Gonzalez et al., 2015; Vlieghe et al., 2005; Dubos et al., 2010; Lee et al., 2008; Maki et al., 2022; Lopes
148 et al., 2023). Genes associated with petal area were significantly enriched for functions related to floral
149 and shoot development (6 genes, $P \sim 0.01$) and reproductive regulation (12 genes, $P \sim 0.02$; Dataset
150 S5). We observed little overlap between SNPs associated with floral and leaf size, consistent with the
151 absence of phenotypic and genetic correlations (**Fig. 1d**). A few genes were associated with both leaf
152 and sepal growth through distinct SNPs, suggesting organ-specific effects of these mutations on
153 general growth regulator activities. Despite moderate phenotypic correlations among floral organs, their
154 genetic determinants showed limited overlap, consistent with low pleiotropy. Thus, floral organ size

155 appears to evolve largely independently, with observed phenotypic covariation more likely reflecting
156 coordinated responses to shared environmental conditions rather than shared genetic architecture.

157 ***Flower-size genes are under a mixed selection pattern in *Arabidopsis thaliana*.***

158 The reduction in floral organ size accompanying the transition to self-fertilization is thought to result
159 from a reallocation of resources no longer required for pollinator attraction toward other fitness-
160 enhancing traits. Under this hypothesis, new mutations reducing flower size should be selectively
161 favored and accumulate in selfing lineages. To test this, we estimated the total effect size and ancestral
162 origin of alleles for all associated SNPs (see Methods). More than three-quarters of derived alleles
163 increased petal area (**Fig. 2a**). This excess of positive effects among derived alleles was unique to petal
164 area (**Fig. 2b**).

165 To assess whether the fixation of derived alleles increasing flower size could result from genetic drift,
166 we first computed genome-wide π_N/π_S , which contrasts the accumulation of non-synonymous (π_N)
167 and synonymous (π_S) diversity to infer the strength and direction of selection. Genes containing SNPs
168 associated with petal area exhibited a skewed π_N/π_S distribution that tended to fall below the genome-
169 wide average (P -value < 0.1), indicating ongoing purifying selection rather than relaxed selection (**Fig.**
170 **2c**). Other floral organ size traits showed no comparable pattern (**Fig. S3**). While most petal candidate
171 genes had π_N/π_S values below the genomic average, three genes, *AT1G36940*, *KIX9* and *FLM*,
172 displayed unusually high values within the top 5% of the genome-wide distribution, suggesting
173 diversifying selection on their protein sequences.

174 We next examined whether derived alleles in petal area genes show evidence of positive selection. At
175 a bi-allelic site, alleles under selection are expected to be associated with longer haplotypes due to
176 incomplete selective sweeps. We estimated haplotype decay around each focal SNP using genome-
177 wide extended haplotype homozygosity (EHH) (Klassmann & Gautier, 2022). The standardized ratio of
178 derived to ancestral haplotype homozygosity (the integrated haplotype homozygosity score, iHS)
179 identifies alleles with unusually long haplotypes for their frequency. Two petal area SNPs exhibited high
180 $|iHS|$ values, resulting from an extended haplotype homozygosity around the derived alleles, a sign
181 of recent positive selection. The first, snp_1_10937139 ($P = 0.013$), increased petal area and likely
182 affects the growth regulator JMJ18. The second, snp_5_870273 ($P = 0.02$), decreased petal area and
183 likely influences AT5G03480, a poorly characterized RNA-binding protein expressed during early petal
184 development. A third SNP, snp_1_28960616 ($P = 0.07$), showed a marginal signal of positive selection
185 and likely affects FLM, a known flowering-time regulator that also influences flower development (**Fig.**
186 **2d**) (Posé et al., 2013). As a control, we repeated the analysis for all floral and leaf traits and quantified
187 the enrichment of trait-associated SNPs within the top 0.05% of the genome-wide iHS distribution
188 (Tsuchimatsu et al., 2020). Leaf size-associated SNPs showed a threefold enrichment, whereas no
189 enrichment was detected for flower size (**Fig. S4**).

190 Together, these results indicate a lack of consistent directional selection acting on the genetic
191 determinants of flower size in *A. thaliana*. Neither purifying nor positive selection acts uniformly across
192 petal size regulators, and no consistent directional trend is apparent. The observed variation likely

193 reflects spatially heterogeneous selection combined with neutral processes, accentuated by the
194 species' small effective population size and strong population structure. We next investigated whether
195 phenotypic and genetic variation in floral organ size arises from spatially heterogeneous selective
196 pressures across *A. thaliana*'s range.

197 ***The influence of flower size on individual's performance is environment dependent***

198 The finding that alleles increasing flower size can be under positive selection in a selfing species
199 challenges the resource allocation theory, which predicts higher fitness associated with reduced
200 investment in floral organ growth in species such as *A. thaliana*. To explore this apparent contradiction,
201 we examined the relationship between petal size variation and plant fitness across populations and
202 experimental environments. We used fitness data from a common garden experiment conducted at
203 three sites, Valencia (Spain), Norwich (United Kingdom), and Halle (Germany) (Wilczek et al., 2014).
204 To maximize overlap between fitness and morphological data, and to remove non-genetic components
205 of petal size variation, we predicted petal size for all 1,135 re-sequenced *A. thaliana* accessions using
206 coefficients from a Bayesian sparse linear model ((Zhou et al., 2013), see Methods). We then compared
207 the predicted petal size with fitness estimates in each experiment. Consistent with expectations of the
208 selfing syndrome, genotypes carrying small-petal alleles generally performed better than those carrying
209 large-petal alleles across most environments (**Fig. 2e and Fig. S5**).

210 However, the relationship between petal size and fitness varied across sites and seasons. In most
211 cases, a quadratic model best described this relationship, with fitness peaking at relatively small petals
212 in Norwich and Valencia, and at intermediate petal sizes in Halle. Under increased environmental
213 stress, such as during summer and autumn in Norwich, the relationship shifted to a negative linear
214 trend, indicating a trade-off between flower size and fitness in harsher conditions. Similarly, other
215 datasets showed a negative association between petal size and fitness, though this trend weakened
216 under extreme stresses (**Fig. S5**).

217 Overall, these results suggest that individual performance is shaped by the genetic propensity to invest
218 in petal growth, but the strength and form of this relationship are environment dependent. Such context
219 dependency likely generates heterogeneous selection regimes on petal size across the geographic
220 range of *A. thaliana*, possibly explaining the lack of unidirectional selection at the species level.

221 ***Habitat suitability influences the selection regime at flower size loci.***

222 Our analyses so far suggest that the fitness costs and benefits of investing in floral growth in *A. thaliana*
223 depend on local environmental conditions. To examine this further, we quantified habitat suitability at
224 the collection sites of the studied populations. We used the MaxEnt (Phillips et al., 2006) to model the
225 species' climatic niche based on bioclimatic variables (Karger et al., 2017), altitude (Danielson & Gesch,
226 2011), and a comprehensive occurrence dataset (Yim et al., 2022). From this model, we extracted the
227 habitat suitability of each studied population (**Fig. 3a, Fig. S6**) and the limiting environmental factors
228 (**Fig. S7**).

229 Across populations, petal area was positively associated with habitat suitability (**Fig. 3b**). However, the
230 relationship was triangular rather than linear: petal size variance increased toward more suitable
231 environments, resembling the previously described relationship between flowering time and habitat
232 suitability (Yim et al., 2022). Quantile regression confirmed this pattern, with the 0.05 and 0.95 quantile
233 slopes differing significantly. This indicates that petal size evolution is more constrained at the climatic
234 margins of the species' range, where environmental stress limits expansion (Yim et al., 2022), while
235 selection on flower size is more relaxed in favourable habitats.

236 To test whether this triangular pattern is general across *A. thaliana* phenotypes, we assessed the
237 relationship between habitat suitability and trait variance in multiple datasets (Atwell et al., 2010;
238 Przybylska et al., 2023). None of the 107 traits in Atwell et al. (2010) showed significant
239 heteroscedasticity, and only whole-plant dry mass in Przybylska et al. (2023) exhibited a similar positive
240 triangular pattern. In our dataset, sepal length, in addition to petal area, displayed both a positive and
241 triangular relationship with habitat suitability. Several other traits showed consistent but homoscedastic
242 responses, including positive correlations for sepal size (this study), flowering time and leaf dry mass
243 (Atwell et al., 2010; Przybylska et al., 2023), and negative correlations for leaf length, rosette diameter,
244 fruit length, and leaf nitrogen content. Together, these results suggest that plant growth and life-history
245 traits broadly align with habitat suitability, and that variance in floral organ size is particularly reduced
246 under less favorable conditions.

247 To explore the evolutionary basis of this pattern, we examined how petal size alleles vary along the
248 habitat suitability gradient. Consistent with the phenotypic trend, large-petal alleles increased in
249 frequency in more suitable environments (**Fig. 3d**). We tested whether this could reflect variation in
250 purifying selection intensity across the species' range. Purifying selection suppresses deleterious
251 mutations, leading to an excess of rare derived alleles and a skewed site frequency spectrum (SFS).
252 We computed the unfolded SFS for genes associated with petal and leaf traits to compare these effects.
253 Low-frequency derived alleles ($0.05 < \text{frequency} < 0.1$) represented about one-fourth of SNPs in both
254 gene sets (**Fig. 3e**). The SFS of petal genes displayed significantly higher skewness in low-suitability
255 habitats, indicating stronger purifying selection at the environmental margin. By contrast, SFS
256 distributions for leaf genes did not differ along the gradient.

257 These results indicate that genes influencing flower size are more constrained in stressful environments
258 and experience relaxed purifying selection in favorable habitats. The accumulation of large-petal alleles
259 in populations from suitable regions therefore likely reflects a release from the fitness penalties imposed
260 by producing large flowers.

261 Discussion

262 Contrary to predictions from resource allocation theory, our findings indicate that petal size variation in
263 *Arabidopsis thaliana* is not governed by a uniform trade-off with fitness. Instead, it reflects a diversity of
264 local evolutionary pressures. Although producing large flowers can be costly, particularly in a species
265 capable of generating up to a thousand flowers (Exposito-Alonso et al., 2018), this constraint is most
266 pronounced in environmentally marginal habitats, where limited resources favor smaller flowers. In

267 contrast, in more favourable environments, such costs appear relaxed, allowing mutations that increase
268 petal size to persist or even spread. Such spatial variation in selection likely explains the greater
269 phenotypic variance in flower size among genotypes from climatically suitable regions.

270 A caveat of our study is that it does not explicitly quantify gene-by-environment ($G \times E$) interactions,
271 which are inherently difficult to assess across the broad environmental gradient occupied by *A. thaliana*.
272 Our measured petal size may not fully represent the phenotype expressed under natural conditions,
273 potentially limiting our ability to infer precise fitness relationships. Nevertheless, previous work shows
274 that $G \times E$ effects on flower size are modest, largely influencing the magnitude rather than the direction
275 of responses to environmental variation (Wiszniewski et al., 2022). Consequently, petal size rankings
276 among genotypes tend to remain stable across environments, as reflected in the significant correlations
277 between our predicted petal area and independent measurements of flower diameter at 17°C ($R = 0.35$,
278 P -value < 0.01), flower diameter at 23°C ($R = 0.46$, P -value < 0.01) (Wiszniewski et al., 2022). Moreover,
279 the selective signatures observed at genes controlling flower size parallel the phenotypic trends,
280 suggesting that functional constraints on small-flower alleles intensify toward the species' climatic limits.

281 In addition to a relaxation of purifying selection, we detected evidence of positive selection acting on
282 derived large-petal alleles, particularly in environmentally favorable regions. This suggests that
283 increased petal size may confer adaptive advantages under certain conditions, or at least when
284 environmental constraints on performance are minimal. The adaptive value of these alleles may relate
285 to the persistence of a low but nonzero rate of cross-fertilisation in *A. thaliana* (Abbott & Gomes, 1989;
286 Hoffmann et al., 2003), potentially favored during post-glacial expansions northward, following
287 secondary contact in the Iberian Peninsula (Fulgione & Hancock, 2018). Until today, the natural
288 variation of selfing rate (s) across a large sample of *A. thaliana* populations has been reported once
289 (Platt et al., 2010). Consistently, we detected a positive correlation between the outcrossing rate ($\log(1-$
290 $s)$) and petal area ($R = 0.25$, P -value = 0.12) and Petal length ($R = 0.32$, P -value = 0.05) (**Fig. S8**).
291 However, given the modest sample size (overlapping genotypes, $N=39$) and multiple-testing context,
292 these relationships should be interpreted cautiously until validated by targeted analyses.

293 Other ecological factors may also promote larger flowers and should be explored in future studies,
294 including historical interactions with floral enemies (Galen, 1999), protection against microbial and viral
295 infections (Vannette, 2020), and the role of flower thermogenesis in pollen germination (van der Kooi
296 et al., 2019). Pleiotropy and linked selection could further contribute to the emergence of large-flower
297 alleles, especially as some petal size-associated variants lie within genes that influence other organs
298 or developmental pathways (Hanemian et al., 2020). While our genetic correlation analyses indicate
299 that such effects are limited, functional validation of causal mutations will be necessary to strictly
300 disentangle developmental coupling from direct adaptive responses.

301 Overall, this study reveals striking geographic variation in investment in reproductive structures,
302 particularly flower size, across the range of *A. thaliana*. Classic theory predicts that selfing lineages
303 experience strong selection for reduced floral display, leading to the fixation of small-flower alleles due
304 to their lower resource costs. Our results support this expectation only under environmentally stressful

305 conditions, where resource limitation likely intensifies selection for smaller flowers. In contrast, in
306 favourable environments, relaxed purifying selection reduces these constraints, enabling the
307 persistence or even evolution of larger-flower variants. Together, these findings illustrate how habitat
308 suitability and resource availability modulate the balance between selection and constraint, shaping the
309 trajectory of phenotypic evolution in selfing species.

310 **Materials and Methods**

311 ***Biological materials***

312 For this study, we subsampled 407 *Arabidopsis thaliana* accessions from (Przybylska et al. 2023 data),
313 originally selected for their extensive geographic and genetic coverage (**Fig. S1, Dataset S1**). All
314 accessions were vernalized for 4 weeks in the dark at 5°C in October 2022. Seedlings were then
315 transferred to soil in individual pots and grown in a glasshouse in semi-controlled conditions in
316 Montpellier, CEFE CNRS, France (maintained around 20°C day and 15°C night with natural light).
317 Accessions were randomly distributed on three tables daily rotated to homogenize the effect of spatial
318 heterogeneity within the glasshouse across the genotypes. We used a single replicate per accession
319 which allowed to maximize the genetic diversity coverage in a restricted experimental area, while
320 genetic relatedness between accessions provided partial genetic replicates. Variance analysis within
321 and between *Arabidopsis thaliana* regional populations (Alonso-Blanco et al., 2016) revealed high
322 heritability for all investigated traits (**Dataset S2**), indicating strong phenotypic similarity among groups.
323 Plants were surveyed daily, and the date of first flower opening, flower collection and plant death were
324 recorded. Three flowers were collected between the 10th and the 25th flower of the main stem from the
325 primary inflorescence. Flowers were collected and preserved in 70% ethanol in individual Eppendorf
326 placed in the dark at room temperature.

327 ***Morphological measurements***

328 Flowers were stained with 0.02% Toluidine Blue for 10 minutes before been dissected and placed in
329 between microscope slides and cover slips with a drop of glycerol (40% v/v). Slides were scanned at a
330 4800 dpi resolution with an Epson perfection v850 pro® scanner. Leaf, petal, sepal, and short and long
331 stamen sizes (length, width, and area when relevant) were measured using Fiji version 2.35 (Schindelin
332 et al., 2012). The blue channel of the images was isolated, and contrast-enhanced with Otsu's local
333 contrast method. Pixels that belong to the organs were counted to estimate organ sizes. Measures were
334 calibrated using the dots per inch (dpi) of the images. The ovule number per flower was counted
335 manually using fluorescence microscopy (Leica DMI4000). Flowering time was calculated as the
336 duration in days from seed sowing to the opening of the first flower. To estimate the genotype leaf sizes,
337 while limiting environmental effects, the fourth leaf of every individual was collected at maturity and
338 immediately scanned at 800dpi resolution. We computed an unbiased predictor of each trait by
339 extracting the residuals of a linear model containing the relevant confounding factors (petal rank along
340 the stem, stem number and a variable describing the spatial distribution in the tables). All traits were
341 normally distributed and did not necessitate transformation (**Fig. S2**). We used our phenotypes in
342 conjunction with published datasets that shared a significant amount of accessions. In particular, we

343 used the 107 phenotypes gathered by (Atwell et al., 2010), containing phenology, microbial resistance,
344 plant morphology and ions content values, and the Aradiv database from (Przybylska et al., 2023) that
345 contains plant functional strategies related traits. We also used datasets reporting fitness in terms of
346 seed production where a majority of our accession set was shared, such as in (Exposito-Alonso et al.,
347 2018; Wilczek et al., 2014), and compared our measurement of flower organ size with other available
348 datasets (Li et al., 2020; Wiszniewski et al., 2022).

349 **Statistical analyses**

350 All statistical analyses were performed with R (version 4.3.2). We performed principal component
351 analyses with *FactoMineR*. Testing for isometry versus allometry was achieved with standard major
352 axis regression, which assumes an error component from both tested variables, with the package *smatr*
353 (Warton et al., 2006). Skewness and deviation from the average were statistically tested in R using a
354 skewness test and a Student's t-test, respectively.

355 **Genome-Wide Association Study (GWAS)**

356 We performed GWAS by testing the association between the above measured organ sizes with publicly
357 available variant calling format (VCF) data from the 1001 genomes project webpage
358 (<https://1001genomes.org/data/GMI-MPI/releases/v3.1/>) (Alonso-Blanco et al., 2016). Although single
359 replicate per accession may reduce detection power by introducing environmental noise, the high
360 heritability of the measured traits and population structure signals associated with phenotypic variation
361 suggest that increasing replicates per accession would have minimal impact on genetic association
362 power (Kang et al., 2008). Moreover, our measurement of petal size correlated significantly with
363 independent flower diameter measurements, i.e. the ranking from small to large flower genotypes
364 remains consistent in between treatments (Wiszniewski et al., 2022) and across experiments
365 (Spearman's Rank Correlation Coefficient ~ 0.46). Only single nucleotide polymorphisms (SNPs) were
366 kept in the analyses, short indels and non-ACTGN were filtered out, as well as SNPs with a minor allele
367 frequency below 5% using PLINK (Purcell et al. 2007). A total of 1,899,962 SNPs passed the filters.
368 GWAS were performed using the GEMMA software v0.98.1 (Zhou & Stephens, 2012), which considers
369 the population structure in a relatedness matrix. The proportion of phenotypic variance explained by the
370 genotypes (PVE, the SNP-based heritability) was estimated for each trait with the *bslmm* function
371 (Bayesian sparse linear mixed model). The P-values distribution of all GWAS followed a uniform
372 distribution and did not necessitate transformation for false discovery rate. Specific quantitative trait loci
373 (QTL) associated with the traits were located in two ways. First we sought major QTL by a genome-
374 wide screening of the association significance, accounting for multiple testing with Bonferroni corrected
375 significance threshold. Second, to consider the potential polygenic architecture of the studied traits, we
376 relaxed the threshold of significance. As there is no consensus on a threshold to detect multiple minor
377 effect variants in a polygenic architecture (also known as multiple disease phenotype), reports of P-
378 values detecting relevant genes range from 10^{-7} to 10^{-4} (Burton et al., 2007; Chen et al., 2021), we
379 used an empirical approach and set the threshold at a P-value where we detected the FLOWERING
380 LOCUS C (FLC) among the flowering time GWAS association peaks (P-value of $5.2 \cdot 10^{-5}$). FLC, known

381 as a key regulator of flowering time, is commonly detected in GWAS in *A. thaliana*, and its variation
382 accounts for more than half the trait's variation (Sasaki et al., 2018). To eliminate genetically linked
383 variants, we filtered out significant SNPs by considering their genomic position
384 (<https://www.arabidopsis.org>), statistical significance, and predicted function using SnpEff (Cingolani
385 MMB, 2012). This process resulted in retaining one focal polymorphism per QTL for subsequent
386 analyses (**Dataset S4**). SNP effect sizes were estimated from the bsImm outputs, as recommended by
387 Zhou et al. (2013) by the formula $\alpha + \beta \cdot \gamma$, with α the small additive effect of each SNP, β the additional
388 effect of some large effect SNPs and γ its posterior inclusion probability estimate. Estimation of the
389 coefficients were further used to predict traits values for all 1,135 sequenced genotypes of the 1001
390 genome project using the predict function from GEMMA (Karasov et al., 2024; Lo et al., 2024).

391 **Population Genetics Analysis**

392 We assessed the regime of selection acting on leaf and flower size QTLs across the species distribution
393 with two metrics. We first computed the π_N/π_S ratio for all *A. thaliana* annotated genes. We calculated
394 the ratio between the genes' nucleotide diversity in non-synonymous sites (π_N – defined as the 0-fold
395 degenerate sites) and the nucleotide diversity in synonymous sites (π_S – defined as 4-fold degenerate
396 sites). The rationale is that π_S gives a measure of neutral genetic variation and serve as a reference
397 for π_N which measures the amount of genetic variation at sites under natural selection control. Thus, a
398 value of π_N/π_S below 1 is indicative of purifying selection, and a value above 1 is indicative of diverging
399 selection. We used the genome-wide distribution as a reference: individual genes' π_N/π_S included in
400 the 0.05 to 0.95 quantile do not deviate from the general purifying trend, Student test and skew test
401 were performed on R to evaluate a set of genes π_N/π_S deviation from average and normality. Full
402 sequences of gene coding regions and general feature format files were downloaded from Phytozome
403 13 (version *Athaliana_447_TAIR10*). Together with the VCF file, they were used to generate the
404 diversity of gene sequences observed among *A. thaliana* genotypes. The R package *RGenetics* was
405 used to discriminate 4-fold from 0-fold degenerate sites, and the *nuc.div* function from *pegas* was used
406 to compute the nucleotide diversity. More details about the methods and scripts used are available at
407 (https://github.com/kevinfrsartori/Whole_Genome_PiNpIS). Second, the extended haplotype
408 homozygosity (EHH) was computed genome-wide to test whether a particular derived variant is
409 subjected to ongoing or partial sweep in populations (Sabeti et al., 2002). When advantageous
410 mutations are selected, the frequency of genetically linked variants also increases, extending the
411 haplotype length around the selected site. This signal perdures until recombination events erase it. EHH
412 measure the decay of a haplotype length upstream and downstream of the focal SNP, which thus
413 depends on the selection's history and strength. Since the patterns of EHH might also be affected by
414 the species' population structure and demographic history (Voight et al., 2006), we compute this
415 estimate genome-wide to assess significance. The R package *rehh* (Gautier and Vitalis Bioinformatics,
416 2012) was used to compute the EHH's test statistic (iHS), defined as the log ratio of the integral of the
417 EHH decay (iHH) around the derived allele divided by the iHH around the ancestral allele. This statistic
418 was then compared to the distribution of iHS for all SNPs of similar allele frequency, and a P-value was
419 reported. The iHS computation relies on the polarization of the SNP dataset, i.e. the ancestral state of

420 each SNP. To differentiate between the derived versus ancestral allele, we used a maximum likelihood
421 method using the focal species allele frequencies together with related species allelic information. We
422 first identified an ortholog gene in *Arabidopsis lyrata* and *Arabidopsis halleri* for most *Arabidopsis*
423 *thaliana* genes by using OrthoFinder (Emms & Kelly, 2015). We kept all *A. thaliana* genes having at
424 least one ortholog and kept only one ortholog per outgroup species based on the genes' distance tree,
425 resulting in a list of 25411 genes. We extracted a sequence covering the coding region plus the cis-
426 and trans-regulatory regions for all trio of genes, and performed a multiple alignment using MAFFT
427 (Kato et al., 2002). We extracted the allelic information of the outgroup species with custom R scripts
428 and ran the program est-sfs (Keightley & Jackson, 2018). It is important to note that the computation of
429 iHS was limited to the presence of ortholog genes, resulting in filtering out few SNP per trait lists
430 **(Dataset S4)**.

431 **Habitat suitability**

432 The *A. thaliana* niche was modelled using the MaxEnt software (Phillips et al., 2006) by reproducing
433 the method from (Yim et al., 2022). Climatic data were downloaded from CHELSA v2.1 ([https://chelsa-](https://chelsa-climate.org)
434 [climate.org](https://chelsa-climate.org)) at a resolution of 30 arc seconds (Karger et al., 2017). The species distribution modelling
435 was performed with the following variable panel: the isothermality (ratio of diurnal variation to annual
436 variation in temperatures, °C), the mean of daily minimum air temperature of the coldest month (°C),
437 the annual range of air temperature (°C), mean daily mean air temperatures of the wettest quarter (°C),
438 mean daily mean air temperatures of the warmest quarter (°C), precipitation seasonality (standard
439 deviation of the monthly precipitation estimates, kg m⁻²), mean monthly precipitation amount of the
440 wettest quarter (kg m⁻² month⁻¹), and mean monthly precipitation amount of the driest quarter (kg m⁻²
441 month⁻¹). Altitude data at a resolution of 30 arc seconds come from the Global Multi-resolution Terrain
442 Elevation Data 2010 (GMTED2010) (Danielson and Gesch 2011) and were downloaded from the USGS
443 website. We used the extensive species occurrence data (N=672) provided by (Yim et al., 2022), which
444 covers the distribution of the species in an evenly spaced manner (one sample per km²) to account for
445 sampling bias. From the MaxEnt model, we estimated the habitat suitability (HS) and the limiting factor
446 (LF) of all *A. thaliana* accessions' collecting sites (**Fig. S6 S7**). The HS metric scales from zero to one,
447 while LF refers to the climatic variable that has the most significant impact on HS when changed while
448 maintaining all others variable constant. The relationships between traits and HS were tested using
449 linear regressions. The apparent heteroscedasticity of several trait-HS relationships was tested by
450 comparing the 0.95 and 0.05 quantile regressions (Peng, 2017) using the R package *quantreg*. Linear
451 regressions were plotted when significant, and quantile regressions were plotted when the test for slope
452 difference (Anova) was significant. All statistical analyses were performed with R (R Core Team, 2020).
453 In order to test for selection regime shift along habitat suitability, we divided the *A. thaliana* accession
454 set into ten equally sized sub-populations. To avoid the influence of over-sampling in some regions of
455 the species distribution, we first pruned the accession list based on genetic relatedness. We used the
456 PLINK software to estimate the number of accessions passing a filtering threshold of genetic
457 relatedness ranging from r=0.1 to r tends to 1. We found that the number of related accessions stops
458 dropping after 0.95 and remains constant while r tends to 1 (**Fig. S9**). We thus excluded 87 highly

459 correlated accessions. HS ranges were obtained by dividing the HS distribution of unrelated accessions
460 into ten quantiles, where each quantile contains about 100 accessions.

461 ***Selection regime along habitat suitability***

462 The HS ranges obtained by splitting the *A. thaliana* accession into 10 populations by habitat suitability
463 did not have a geographic and historic meaning for the species, which prevents from computing most
464 classical population genetic metrics. We tested for changes in selection pressure by relying on allele
465 frequency only. We used the sequencing dataset presented above and reduced the minimum allele
466 frequency threshold to a minimum allele count of 1. Only the SNPs having ancestry information located
467 in the gene body were kept, and we computed the unfolded site frequency spectrum for leaf and flower
468 organs associated genes by reporting the allele counts for ranges of allele frequency of 0.05.

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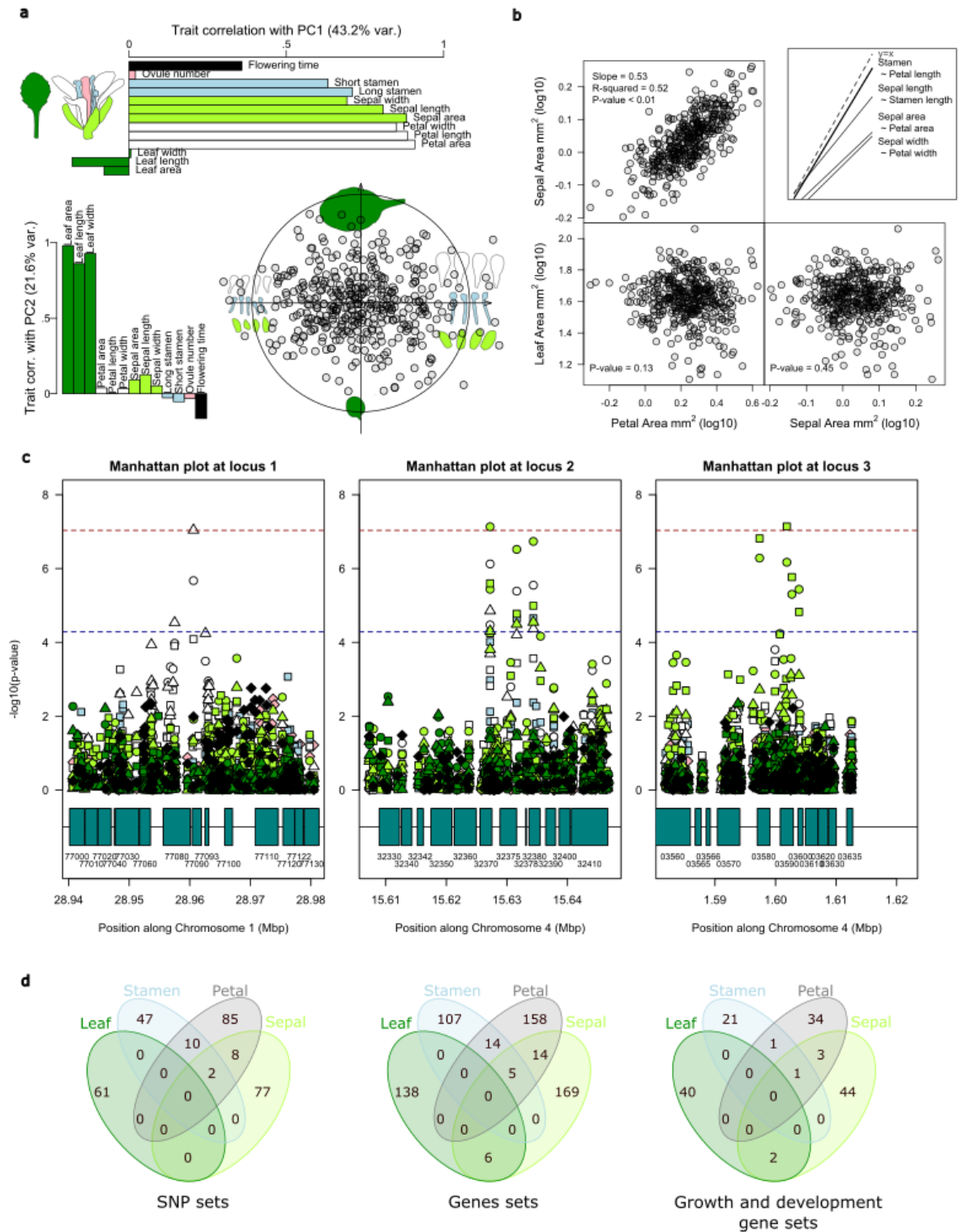
478 **References**

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482 **Figures and Tables**

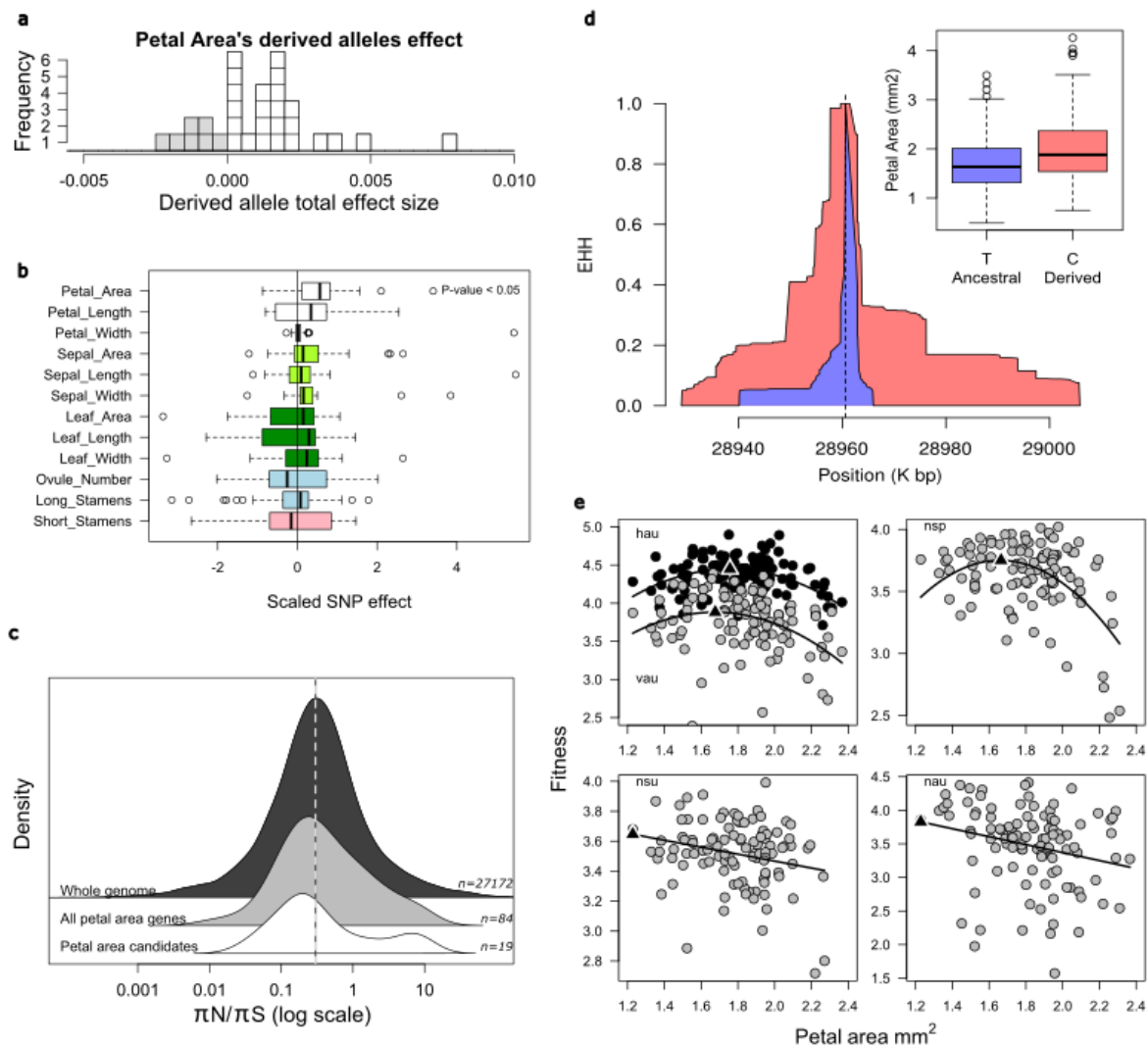


484 **Figure 1: Extensive flower size variation in *A. thaliana* results from low pleiotropic**
485 **mutations.**

486 **a)** A Principal component analysis (PCA) was performed on flower organs and leaf size
487 excluding flowering time that was correlated post hoc and represented as a supplementary
488 variable. Bar plots show trait correlations with principal components (PC), with colours
489 indicating different organs as shown on the top right scheme: dark green, leaf; light green,
490 sepal; white, petal; blue, stamen; pink, ovules. The biplot displays accession coordinates along
491 the first two principal components (PC), with organ silhouettes representing extreme organ
492 size values at the same scale; **b)** Pairwise relationships between organ sizes. Coefficients
493 were obtained from standard major axis linear regressions. The top right plot illustrates the
494 slope test for organ isometry (null hypothesis: slope = 1). **c)** Manhattan plots displaying the
495 genomic regions surrounding the three significant associations peaks (based on Bonferroni
496 threshold) detected for all studied traits. Dot colours indicate associations with the size of
497 different organs as depicted in a), while shapes correspond to the parameter analyzed
498 (triangles indicate association with area, circles with width or short, and squares with length
499 or long). The Red dashed line represents the Bonferroni threshold, while the blue dashed line
500 represents the polygenic threshold. Blue rectangles materialize the annotated gene code
501 numbers (following “ATxG” in TAIR nomenclature, with x being the chromosome number). **d)**
502 Venn diagrams illustrating the lack of overlap between SNPs, genes and candidate genes
503 associated with leaf and flower organs’ size variation.

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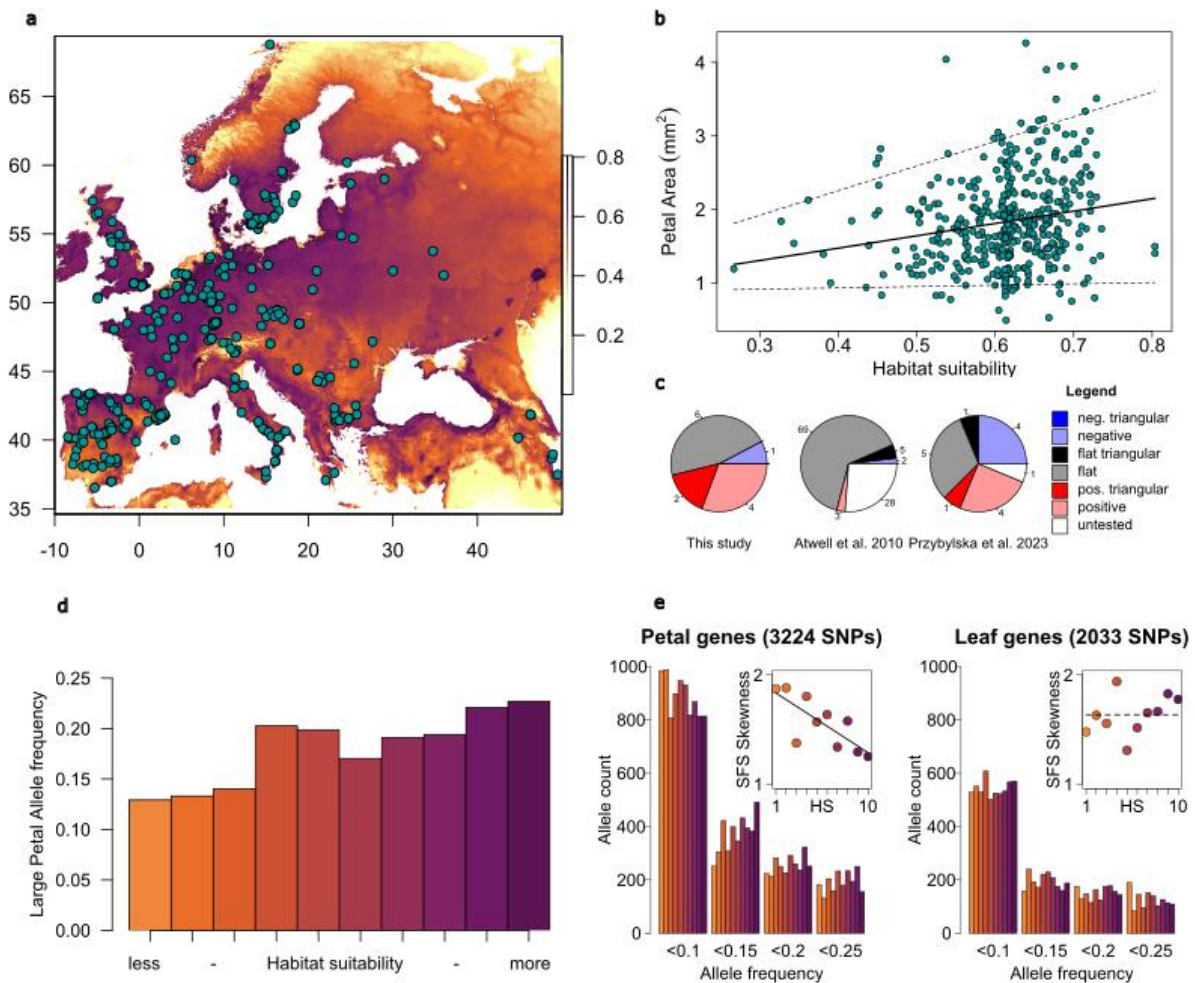
507 **Figure 2: A variable flower size–seed production trade-off explains a mixed selection**
 508 **pattern on flower size in *A. thaliana*.**

509 **a)** Distribution of derived alleles' total effect size from each peak of association identified in
 510 GWAS for petal area. Each square represents a SNP's derived allele, white and grey
 511 indicating a positive and negative effect, respectively. **b)** Distribution of derived alleles' total
 512 effect size from each peak of association identified in GWAS for each trait investigated. A p-
 513 value is given for trait distributions significantly different from 0 (vertical line). **c)** $\pi N/\pi S$
 514 distribution across all *Arabidopsis thaliana* genes, among the petal area genes and a priori
 515 candidate genes identified in the GWAS for petal area. The dashed line represents the whole
 516 genome average. **d)** Extended haplotype homozygosity around FLM lead SNP position. The
 517 red and blue integrals represent the degree of haplotype identity shared by individuals carrying
 518 the C (derived) and T (ancestral) alleles, respectively, as a function of distance to the lead
 519 SNP (dashed line). **e)** Relationship between the genetic variance of petal area and seed

520 production across different experimental fields; hau, Halle autumn; vau, Valencia autumn; nsp,
521 Norwich spring; nsu, Norwich summer; nau, and Norwich autumn. Triangles represent the
522 modelled petal size that maximizes seed production, black lines represent the best fit (anova)
523 between a linear and a quadratic regressions, and blue shadows represent the confidence
524 intervals of the regressions.

525

526



527

528 **Figure 3: Habitat suitability shapes patterns of phenotype diversification across the**
 529 **geographical range of *A. thaliana*.**

530 **a)** Map of the studied European accessions. The orange to purple scale represents the
 531 modelled habitat suitability from less to more suitable, and the blue dots represent accessions'
 532 collecting sites. **b)** Correlation between petal area and habitat suitability. Lines represent
 533 significant regressions: plain line, linear model; dashed lines, 0.05 and 0.95 quantile
 534 regressions. **c)** Large petal alleles' frequency along habitat suitability clusters. **d)** Unfolded site
 535 frequency spectrum (SFS) among SNPs located in petal and leaf gene sets. The colours of
 536 the bars correspond to the habitat suitability classes. Inserts represent the relationship
 537 between the SFS's shape and habitat suitability, where a plain line depicts significant and
 538 dashed line insignificant correlation.

539