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Boumghar, P Saviuc, A Zegnhoun, L Filleul, et al.

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Short-Term Effects of Air Pollution on Mortality in Nine French Cities: A Quantitative Summary

ALAIN LE TERTRE PHILIPPE QUÉNEL DANIEL EILSTEIN · SYLVIA MEDINA HELENE PROUVOST I.AURENCE PASCAL Environmental Health Department lnstitut de Veille Sanitaire (lnVS) Paris, France AZZEDINE BOUMGHAR Regional Observatory of Health Paris, France PHILIPPE SAVIUC ABDELKRIM ZEGHNOUN LAURENT FILLEUL Environmental Health Department Institut de Veille Sanitaire (lnVS) Paris, France CHRISTOPHE DECLERCQ Regional Observatory of Health Lille, France **SYLVIE CASSADOU** CORINNE tE GOASTER Environmental Health Department lnstitut de Veille Sanitaire (lnVS) Paris, France

> ABSTRACT: Between 1990 and 1995, 9 French cities provided data on daily air pollution, total mortality, cardiovascular mortality, and respiratory mortality. Personnel in individual cities performed Poisson regressions, controlling for trends in seasons, calendar effects, influenza epidemics, temperature, and humidity, to assess the short-term effects of air pollution. The authors describe results obtained from the quantitative pooling of these local analyses. When no heterogeneity could be detected, a fixed-effect model was used; otherwise, a random-effect model was used. Significant and positive associations were found between total daily deaths in these cities and the 4 air pollution indicators studied: (1) Black Smoke, (2) sulfur dioxide, (3) nitrogen dioxide, and (4) ozone. A 50-µg/m³ increase in Black Smoke (24hr), sulfur dioxide (24hr), nitrogen dioxide (24hr), or ozone (B hr) was associated with increases in total mortality of 2.9% (95% confidence interval [CI]) = 1.3, 4.4), 3.6% (95% Cl = 2.1, 5.2), 3.8% (95% Cl = 2.0, 5.5), and 2.7% (95% Cl = 1.3, 4.1), respectively. Similar results were obtained for cardiovascular mortality. Except for sulfur dioxide, positive-but not significant-associations were found with respiratory mortality. The internal consistency among the cities studied, as well as consistency with previously published results, favors a causal interpretation of these associations.

> <Key words: air pollution, Black Smoke, cardiovascular mortality, nitrogen dioxide, ozone, respiratory mortality, short-term effects, sulfur dioxide, total mortality>

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THE RESULTS OF RECENT EPIDEMIOLOCICAL STUDIES have shown that ambient air pollution can adversely affect human health, even at levels close to or lower than current national standards.¹ Particulate matter is the pollutant most consistently associated with mortality; however, the lack of such an association for other pollutants, especially sulfur dioxide $(SO₂)$, has been questioned. In particular, the independent mortality effect of particulate matter has been shown in the United States² but not in Europe, where the effect of SO_2 has not been modified by particle levels.³ Ozone (O_3) has also been associated with both total and cause-specific mortality; respiratory and cardiovascular systems are mainly affected by O_3 .⁴In a few studies, investigators have examined the effects of nitrogen dioxide ($NO₂$) on mortality and have suggested the existence of an adverse effect; however, the independence of this effect remains unclear.⁵

The association between air pollution and daily mortality has been studied in Europe, Iargely within the Air Pollution and Health: A European Approach (APHEA) project. In this study, investigators found consistent results in large Western European cities.³⁻⁵ The effects of air pollution have been well established in European metropolitan areas; however, few European studies have been conducted in relatively small urban areas. The investigators of the French 9-cities project, conducted between 1990 and 1995, attempted to quantify the short-term effects of air pollution on mortality in medium-sized cities. The investigators chose 9 medium-sized cities in which air pollution measurements had been made during the past few years. ln addition, the investigators' choices were influenced by (1) the daily number of health outcomes, which vary in accordance with the size of a city; and (2) their desire to select cities that contrasted with each other with respect to climate, air pollution sources, and socioeconomic status of the population.

Materials and Method

ln Table 1 are shown population sizes for the cities studied (i.e., B medium-sized cities and the Paris region). Also shown are daily mean deaths for (1) total mortality (based on the lnternational Classification of Diseases, 9th revision $[ICD9]$; < 800), (2) cardiovascular mortality (ICD9 390-459); and (3) respiratory mortality (ICD9 460-519). The cities have a combined population of nearly 11 million, including 6 million in the Paris metropolitan area. Mortality data were taken from national statistics. The daily number of deaths was lowest in Le Havre, the smallest city, and highest in Paris. Cardiovascular mortality ranged from 30% of total mortality in Rouen to 37% in Marseille. The daily number of deaths from respiratory causes was close to 1 in all cities, except Paris, and respiratory causes represent approximately 7% of the total deaths.

The daily means for meteorological variables for the study period are given in Table 2. The average yearly low temperature ranged from 6.6 °C in Rouen to 10.9 \degree C in Marseille. Marseille reported the lowest humidity level and Le Havre reported the highest. Air pollution monitoring network stations established in each city provided hourly or daily air pollutant measurements. The pollutants studied were those generally measured in France: $SO₂$ was measured by ultraviolet fluorescence (except in Bordeaux, where it was measured by strong acidity); $NO₂$ was measured by chemiluminescence; O_3 was measured by ultraviolet absorption; and Black Smoke was measured by reflectometry. Daily 24 hr means were used for all pollutants, except O_3 for which B-hr maximum values were used.

We encountered a problem in collecting data from the monitoring stations during the study period. For some cities, data were available for only 1 pollutant, and for other cities, data on pollutants were unavailable until recent years. Three cities (i.e., Le Havre, Paris, and

 $tICD9 = 390-459.$

 \pm ICD9 = 460-519.

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Rouen) provided measurements for all 4 pollutants. Only $SO₂$ data were available for the city of Lille. Marseille had data for Black Smoke and $SO₂$ only, and Toulouse reported for $NO₂$ and $O₃$ only. Descriptive statistics for each pollutant, and the time period during which data were available, are provided in Table 3.

To ensure between-city comparability of pollutant exposure, only data from urban background stations were used, except for O_3 for which suburban station data were used. France has a nationally standardized classification system for air quality'monitoring stations, and there are strict rules for station operation.⁶ The correlation between daily measurements at different stations in the same city ranged from 0.34 to 0.84 for SO_2 , 0.55 to 0.84 for Black Smoke, 0.52 to 0.83 for $NO₂$, and > 0.95 for O_3 . Six cities could provide measurements for O_3 ; however, 3 cities used only 1 station each, and 2 cities used 2 stations each. For the other pollutants, all cities collected informatiôn from at least 2 stations, except for Strasbourg, where only 1 station was available for each pollutant. The range of median pollutant levels did not differ greatly among cities, inasmuch as the maximum differential was only 42 μ g/m³ for median levels of O₃ (Table 3).

We analyzed the daily data from the 9 cities by Poisson time-series regressions. We fit the data with autore-

gressive models that allowed for overdispersion. Adhering to APHEA methodology,^{7} we controlled for possible confounders, including long-term trends, seasons, days of the week, holidays, influenza epidemics, temperature, and humidity. lnstead of using a parametric approach that controlled for confounders (as was done in APHEA), we used generalized additive models⁸ to ensure a nonlinear relationship between confounders and mortality. As part of the diagnostic testing of the models, we tested the residuals for autocorrelation, applied Akaïke's information criteria, and plotted the residuals.

Statisticians from each monitoring station performed Poisson's time-series regression with their station's own data. To ensure comparability between the modeling processes, we subjected each model to quality control by the coordinating center, in collaboration with the statisticians from every center. Each center determined the best lag time for each pollutant (i.e., 1-5 days or cumulatively Ii.e., the mean over several previous daysl). Finally, for the meta-analysis, each center fitted a linear relationship between the average of lags 0-1 (day) of the pollutant and mortality. This approach avoided potential bias that could have resulted from use of the most significant lags in each city. We, therefore, based our choice of the Iinear form on the results from the local analyses.

To obtain a quantitative summary of the results across the centers, we applied meta-analysis methods to arrive at a pooled regression coefficient as a weighted mean of the city-specific regression coefficient. The weights are inversely proportional to the city variances. We then determined the weights, assuming a fixed-effect model when a chi-square test failed to detect heterogeneity at the sensitivity level of alpha $= 20\%$. If the assumption of homogeneity was rejected, we adopted a random-effect model in which the between-cities variance was added to each estimated local variance, giving more similar weights—but also a larger variance. This variance inflation reflects greater uncertainty on the summary estimate when local results are heterogeneous. We used the moment method of Der Simonian and Laird⁹ to estimate the between-cities variance from the data. Civen

that Paris represented half of the population of all the cities, we conducted a sensitivity analysis to determine whether changes occurred in the fixed estimates when Paris data were excluded.

Results

Each city's relative risk (RR) of mortality associated with a $50-\mu g/m^3$ increase in the average pollutant level for 0- and 1-day lags is shown in Tables 4-7.

 $SO₂$. The average total mortality for each city, as well as pooled fixed-effect (with and without Paris included) and random-effect RRs (and their 95% confidence intervals $[CIs]$) associated with a 50- μ g/m³ increase in the average level of $SO₂$ after 0- and 1-day lags are shown in Figure 1. All local estimates were positive (i.e., RR > 1), although they reached a nominal significance level of 5% in the 4 cities of Bordeaux, Lille, Lyon, and Paris. The $SO₂$ results were heterogeneous across cities (p [for heterogeneity] = .13) In attempting to explain this heterogeneity, we found that separating the Bordeaux data from that of the other cities resulted in more homogeneous subgroups. A possible explanation is that Bordeaux was the only city that used strong acidity as a proxy for SO_2 . All other cities presented a homogeneous estimated RR. The random joint effect (RR $=$ 1.041) was greater than the fixed effect (RR = 1.036),

but its 95% Cl was also wider for the random joint effect $(95\% \text{ Cl} = 1.018, 1.064 \text{ for random and } 1.021$. 1,052 for fixed). We also found heterogeneity in results for $SO₂$ relative to cardiovascular deaths in Bordeaux. Inasmuch as CIs were wide for $SO₂$ relative to respiratory mortality, no significant heterogeneity was foundeven with the higher RR in Bordeaux. The exclusion of Paris gave a similar pooled RR as described above for total mortality, but there were a higher RR and wider Cls for cardiovascular and respiratory mortality ($p = .05$). A 50-µg/m³ increase in SO₂ (Table 8) was associated with an increase of 3.6% in total mortality, 5.3% in cardiovascular mortality, and 5.6% in respiratory mortality.

Black Smoke. Figure 2 shows our findings for cardiovascular mortality for individual, as well as pooled fixed-effect, RRs (with and without Paris), and their 95% Cls, associated with a $50-\mu g/m^3$ increase in the average level of Black Smoke after 0- and 1-day lags. The Black Smoke results were consistent across cities (p [for heterogeneity] = .75). All local estimates were positive, but none were significant. Under the fixed-effect model, a 50 -µg/m³ increase in the daily 0- to 1-day average BIack Smoke levels was associated with an increase in total mortality equal to 2.9% (95% Cl = 1.3, 4.4). The increase in daily cardiovascular deaths was 3.1% (95%) $Cl = 0.5$, 5.5). The pooled estimate (Table 8) was positive ($RR = 1.027$) for respiratory mortality, but it was not

City	Total mortality		Cardiovascular mortality		Respiratory mortality	
	RR	95% CI	RR	95% CI	RR	95% CI
Bordeaux	1.198	1.054, 1.361	1.353	1.098, 1.666	1.270	0.799, 2.019
Le Havre	1.012	0.983, 1.041	1.049	0.996, 1.104	1.061	0.956, 1.178
Lille	1.045	1.010, 1.081	1.059	1.002, 1.120	1.061	0.947, 1.190
Lyon	1.097	1.023, 1.176	1.064	0.945, 1.199	1.092	0.859, 1.389
Marseille	1.058	0.996, 1.125	1.149	1.041, 1.268	1.084	0.879, 1.335
Paris	1.035	1.008, 1.063	1.026	0.980, 1.075	1.028	0.940, 1.125
Rouen	1.015	0.948, 1.087	1.013	0.893, 1.148	1.157	0.904, 1.482
Strasbourg Toulouse	1.040	0.948, 1.141	1.008	0.864, 1.176	0.944	0.670, 1.329

Table S.-Estimated Relative Risks (RRs) of Mortality and 95% Confidence lntervals (Cts) Associated with a 50-µg/m³ Increase in Black Smoke Levels across 9 French Cities Studied

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Table 6.-Estimated Relative Risks (RRs) of Mortality and 95% Confidence lntervals (CIs) Associated with a 50-µg/m³ Increase in Nitrogen Dioxide Levels across 9 French Cities Studied

	Total mortality		Cardiovascular mortality		Respiratory mortality	
City	RR	95% CI	RR	95% CI	RR	95% CI
Bordeaux						
Le Havre	1.021	0.954, 1.093	1.078	0.955, 1.217	1.223	0.949, 1.575
Lille						
Lyon	1.005	0.922, 1.096	1.049	0.905, 1.215	0.996	0.743, 1.334
Marseille						
Paris	1.042	1.022, 1.062	1.046	1.010, 1.083	1.022	0.953, 1.096
Rouen	1.044	0.970, 1.123	1.159	1.019, 1.319	1.054	0.813, 1.367
Strasbourg	1.000	0.933, 1.072	0.934	0.828, 1.054	1.066	0.822, 1.383
Toulouse	1.076	0.957, 1.210	1.023	0.854, 1.226	1.286	0.821, 2.015

Table 7.-Estimated Relative Risks (RRs) of Mortality and 95% Confidence lntervals (Cls) Associated with a 50-µg/m³ Increase in Ozone Levels across 9 French Cities Studied

Fig. 1. Estimated individual and pooled relative risks (RRs) of total daily number of deaths and 95% confidence intervals associated with a 50-µg/m³ daily increase (i.e., 24 hr) in levels of sulfur dioxide.

Table B.-Estimated Pooled Relative Risks (RRs), with and without Paris, of Mortality and 95% Confidence Intervals (CIs) Associated with a 50-µg/m³ Increase in Levels of Pollutants across 9 French Cities Studied

Pollutant	Total mortality		Cardiovascular mortality		Respiratory mortality	
	RR	95% CI	RR	95% CI	RR	95% CI
Black Smoke	$1.029*$	1.013, 1.044	1.031	1.005, 1.055	1.027	0.974, 1.083
(24 hr)	1.046+	1.019, 1.074	1.055	1.009, 1.104	1.111	1.013, 1.219
SO ₂ (24 hr)	1.036	1.021, 1.052	1.053	1.026, 1.080	1.056	1.002, 1.113
	$-1,036$	1.017, 1.055	1.065	1.032, 1.098	1.071	1.004, 1.143
NO ₂ (24 hr)	1.038	1.020, 1.055	1.046	1.015, 1.078	1.040	0.978, 1.106
	1.023	0.988, 1.060	1.046	0.984, 1.111	1.103	0.971, 1.253
O_3 (8 hr)	1.027	1.013, 1.041	1.024	0.997, 1.051	1.008	0.952, 1.062
	1.033	1.013, 1.053	1.026	0.990, 1.064	1.035	0.962, 1.114

Notes: SO_2 = sulfur dioxide, NO₂ = nitrogen dioxide, and O_3 = 0zone. *RR calculated under a fixed-effects model, with Paris. tRR calculated under a fixed-effects model, without Paris.

significant (95% Cl = 0.974 , 1.083). Exclusion of Paris gave a higher and significant RR $(= 1.111)$ in relation to respiratory deaths $(95\% \text{ Cl} = 1.013, 1.219)$.

NO2. Figure 3 shows the statistics on respiratory mortality for individual and pooled fixed-effect RRs (with and without Paris), and their 95% Cls, associated with a 50-µg/m³ increase in the average level of $NO₂$ after a 0- or 1-day lag. The $NO₂$ results were consistent across cities (p [for heterogeneity] = .73). All local estimates were positive, but none were significant. Under the fixed-effect model, a 50- μ g/m³ increase in the daily 0to 1-day average $NO₂$ levels was associated with a 4.0% increase (95% Cl = -2.2 , 10.6) in daily respiratory deaths. Effects on total mortality and cardiovascular mortality were 3.8% and 4.6% , respectively, and were significant at a nominal level of 5%. Exclusion of Paris gave systematically wider and nonsignificant Cls, ^a smaller RR for total mortality, a similar RR for cardiovascular mortality, and a higher RR for respiratory mortality (Table 8).

 \mathbf{O}_3 . Pooled-effect estimates of O_3 were similar (in the 6 cities for which data were available) for total and cardiovascular mortality (Table 8). All local estimates were positive, although none were statistically significant, except total mortality for Paris. Pooled estimates were not sensitive to the inclusion of Paris. A $50-\mu g/m^3$

increase in O_3 was associated with a 2.7% increase in total mortality and a 2.4% increase in cardiovascular mortality. Results for respiratory mortality were less conclusive than for total mortality, as we found a negative effect in 2 cities, a null effect in 1, and a positive effect in 3 cities. The CIs were very large in all cities. The RR associated with a 50-µg/m³ increase in O_3 was 1.008 (95% Cl = 0.952, 1 .062). Exclusion of Paris gave a higher risk ($RR = 1.035$), but wider CIs also resulted.

Discussion

Our combined analysis summarizes the results from 9 French cities on the short-term effects on mortality of 4 indicators of air pollution. We found that all the pol-Iutants studied had a clear and significant effect on total mortality, with a 50- μ g/m³ increase in pollutants being associated with an increase of $2.7\% - 3.8\%$ in total daily deaths. In the European meta-analysis (i.e., APHEA project), investigators estimated that an increase of 50 μ g/m³ in air pollution levels was associated with an increase in total mortality of $1.3\% - 2.9\%$.^{3,5} Results for cardiovascular mortality showed a significant RR for Black Smoke, SO_2 , and NO_2 , ranging from 3.1% to 5.6%, with a 50- μ g/m³ increase in these pollutants. Nonsignificant, but positive, associations were found with increases in O_3 . Results for respiratory mortality were less conclusive, as we found no significant relationship between increased mortality and pollutants, except with SO₂. Moreover, coefficients were closer to 0 and very sensitive to the inclusion of Paris. Kinney and Hozkaynak¹⁰ also found no significant relationship

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between increased O_3 and respiratory mortality. They suggested that the small number of deaths from respiratory causes might have limited their ability to detect minor pollution associations.

It should be noted that certain problems frequently encountered in meta-analysis do not apply to the present study. There was no selection bias because the participating cities were not selected on the basis of the results of short-term analyses, but, rather, on their ability to provide data. No cities or results were later excluded. No important differences were seen in health endpoints, exposure data, or the analysis procedures used. We set rules (see Method section) for exposures and confounding factor measurements. ln each city, statisticians conducted sensitivity analyses by including different temperature lags, excluding daily extremes in temperature or pollutant levels, and by the use of robust regression.

The use of meta-analysis for the study of air pollution effects is commonly criticized because the combination of RRs, chosen at different lags, is considered, and the end result is that the impact of air pollution findings is maximized. We attempted to avoid this pitfall by choosing the cumulative 0- to 1-day lags a priori based on previous studies. However, defining a priori a fixed lag for each city can induce heterogeneity in the results, as the effect of air pollution may differ by time between cities, and may be affected by the overall (poor) health of a population or the extent of exposure to air pollution.

No between-city heterogeneity was found in the results for total, cardiovascular, or respiratory mortality, except with SO_2 . A plausible explanation is that Bordeaux was the only city that measured strong acidity as

 \log roxy for SO₂. Coefficients from other cities gave mogeneous results, as did a pooled fixed estimate se to the random estimate, including Bordeaux. reover, an increase of 50 μ g/m³ in strong acidity in rdeaux represented the full range of this pollutant's ribution. There was no concurrent measurement of ing acidity and SO_2 in Bordeaux; however, the annumean SO₂ in the years that followed the study period $\overline{\text{c}}$ one-third higher than the annual means for strong rity. Applying this conversion factor in our study no noticeable change in the overall effect, d, perhaps strong acidity could better capture expoad to homogeneity in the results. On the other e to aerosol acidity. Gwynn et al.¹¹ reported a speassociation of sulfates (SO_4^{2-}) and hydrogen ions with total mortality or hospital admissions, and restory mortality or admissions, but no associations und for circulatory mortality or admissions. Sulare associated with respiratory disease and mory.¹² Factor analysis has revealed that both combusrelated pollutants and secondary aerosols (sulfates) associated with cardiovascular mortality.¹³ When compared cities that had a relative increase in interrtile range-rather than a fixed range-we obtained neous results. Results from individual cities had s, as daily death counts were low, except in is. Detection of possible heterogeneity is related to monitoring capabilities of the individual cities in larger CIs in each city are more susceptible to over-To compensate for this lack of capabilities, we perred graphical analyses by ordering local RRs in ns of different variable levels. These variables uded medians of pollutants or other pollutants studduration of the series, number of stations used, '
I : mean mortality level, and mean temperature. No r relationship was found between any of the aforertioned variables.

iven that a few cities could provide data for all polnts, and some provided data for only a brief period, were unable to explore in detail possible effect such as analyses of local results from a 2-pol- π

air pollution data were not reliable in France at the ng of the 1990s, the situation has now changed w passed in December 1996 requires that policies erning the control of air pollution take into account mpact on public health. The law stipulates that an , uality program is to be set up in each region, on the s of an assessment of emission sources, emission and health impacts. This new policy climate Ild provide an opportunity for more complete polan analysis in the future.

sitivity analysis was conducted on the potential
f the inclusion of the Paris results. When the is results were included, RRs tended to be driven by '--
'e results and CIs were wider. Yet, there was no great repancy between pooled fixed effects with and out Paris for total and daily cardiovascular deaths. piratory analyses were less conclusive than were
liovascular analyses—results from city to city were homogeneous and reflected wider CIs. The RRs for

respiratory mortality were more influenced by the inclusion of Paris data.

As is the case in all ecological studies, there is some misclassification of exposure, to a degree that could vary across cities. Personal exposure to outdoor pollutants is poorly estimated by the fixed monitoring stations because there is geographical heterogeneity, and there are indoor sources of pollution. Nevertheless, in time-series studies, days, rather than persons, are the units of observation. lf this is an accurate approach, then personal exposure is poorly correlated to monitoring stations. Some investigators have shown that day-today variations in personal exposure are correlated strongly with daily variations in air pollution levels measured by monitoring stations. $14-17$

The biological mechanisms underlying the association between daily variation of exposure to pollutant levels and daily mortality are still debated. Particles seem to be the pollutant for which recent biological research indicates a possible causal effect. Recent findings tend to relate particulate pollution to increased plasma viscosity,¹⁸ increased risk of elevated heart rate,¹⁹ and electrocardiographic changes in dogs.²⁰

 $SO₂$ is a respiratory irritant and bronchoconstrictor that can induce coughing and wheezing. Civen that $SO₂$ is highly correlated with particles in some U.S. cities, Schwartz et al. 21 postulated that this association could reflect the effect of particles. However, in the APHEA project, the effect of $SO₂$ was not modified by particle levels.³⁻⁵ Although the mechanism behind acidic aerosol health effect associations is not yet known, several have been suggested. Winchester²² argued that acid-coated particles could lead to the solubilization of metals in the lung. Seaton et al.²³ hypothesized that acidic, ultrafine aerosols provoke alveolar inflammation.

It is well documented that experimental exposure to high levels of $NO₂$ can alter pulmonary function.²⁴ In epidemiological studies, researchers found an association between current levels of $NO₂$ and mortality.^{5,25,26} However, this association was not as strong when particles were included in the model. These results imply that $NO₂$ may serve as a proxy variable for vehiclederived pollutants.

At current levels, O_3 has had acute pulmonary effects on susceptible populations in the cities we studied.²⁷ Even at low levels, O_3 can induce inflammation. In epidemiological studies, findings on the adverse effects of $O₃$ have consistently shown an association with total mortality and with emergency hospitalizations for respiratory disease-mainly asthma.^{28,29}

In conclusion, the coherence of results across various cities and studies, especially for Black Smoke and O_{3} along with what is known of possible mechanismsreinforces our argument that the associations found in this study were most likely causal.

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Requests for reprints should be sent to Dr. Alain Le Tertre, lnVS, 12 rue du val d'osne, 99410 Saint Maurice, France. E-mail: a.tertre@invs.sante.fr

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