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Spatio-temporal dynamics of attacks around deaths of wolves: A statistical assessment of lethal control efficiency in France

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18 Abstract

19 Evaluating the efficiency of lethal control of large carnivores such as wolves to reduce attacks on

20 livestock is important given the controversy surrounding this measure. We used retrospective data

21 over 10 years and an intra-site comparison approach to evaluate the effects of lethal control on the

22 distribution of attack intensities in the French Alps. We analyzed 278 legal killings of wolves between

23 2011 and 2020 and the 6110 attacks that occurred during a period of ± 90 days and within 10 km

around these lethal removals. This large number of attacks allowed us to perform an original

25 framework that combined both continuous spatial and temporal scales through 3D kernel

26 estimation. We also controlled the analysis for livestock presence, and explored different analysis

27 subsets of removals in relation to their locations, dates and proximity to other removals. This

28 statistical method provided an efficient visualization of attack intensity spatio-temporal distribution

- 29 before and after removals. A decrease of the intensity of attacks was the most common result after
- 30 the lethal removals of wolves. However, this outcome was not systematic for all subsets and
- 31 depended on the scale of the analysis. In addition, attacks tended to persist after removals while
- 32 showing frequent interruptions in time after but also before removals. We also observed localized
- 33 positive trends of attack intensities at varying distances from removals after they occurred. To
- 34 summarize, our results showed that considering the scale of the analysis is crucial and that effects
- 35 should be analyzed separately for each local context. As a next step, we recommend to move
- 36 forward from patterns to mechanisms by linking the effects of lethal control on wolves to their

37 effects on attacks through analysis of fine-scaled data on wolves and livestock.

38 *Keywords:* before-after study design, *Canis lupus*, culling, depredation, French Alps, kernel density

39 estimation, removal, sheep

40 **1. Introduction**

- 41 Lethal control of large carnivores aims to reduce damages on livestock by decreasing the encounter
- 42 risk with livestock or by removing individuals inclined to predation on livestock (Lennox et al., 2018).
- 43 The side effects of these management practices are often claimed to undermine the efficiency of
- 44 lethal control or even to create more damages (Elbroch and Treves, 2023). In the specific case of
- 45 social canids like wolves, studies have mostly pointed out additive human-caused breeder loss as a
- 46 risk for increasing damages on livestock (Jęodrzejewska et al., 1996; Žunna et al., 2023). Social
- 47 instability can indeed cause multiple reproduction (Ausband et al., 2017) or pack disruption (Cassidy
- 48 et al., 2023), although their effects on damages have not been assessed.
- 49 Implementing controlled experiments to evaluate the efficiency of lethal control of large carnivores
- 50 is extremely difficult, considering the involved large spatial and temporal scales. To date, no
- 51 experimental study has been conducted for this matter (Treves et al., 2016, 2019). Only Allen (2014)
- 52 conducted a controlled experiment to evaluate dingo's baiting efficiency but using the number of
- 53 lactation failures as a proxy for calves lost to predation instead of direct cases of predation.
- 54 Also, retrospective studies have been conducted to answer the question of lethal control efficiency.
- 55 The data they used were not collected within a scientific frame but generally through management
- 56 (*e.g.* damages from compensation schemes). Most of retrospective studies on social canids (gray
- 57 wolves *Canis lupus*, dingoes *Dingo dingo*, coyotes *Canis latrans*) were correlative and tested for a
- relationship between the number of livestock damages and the number of killed predators (Šuba et
- al., 2023; Fernández-Gil et al., 2016; DeCesare et al., 2018; Conner et al., 1998; Allen, 2015; Kutal et
- al., 2023). However, this approach is not suited to infer causality and is subject to high risks of
- 61 confounding factors and to analytical variability (Gould et al. 2023). For example, three correlative
- 62 studies which used the same dataset concluded differently on the effect of lethal removals of wolves
- on livestock damages (Kompaniyets and Evans, 2017; Poudyal et al., 2016; Wielgus and Peebles,
- 64 2014).
- 65 A smaller number of retrospective studies used a comparative approach. Harper et al. (2008),
- 66 Bradley et al. (2015), Santiago-Avila et al. (2018) and Wagner and Conover (1999) compared
- 67 situations where lethal control was used and situations where it was not, *i.e.* made inter-site
- 68 comparison. Bjorge and Gunson (1985) and Blejwas et al. (2002) compared the situation before and
- 69 after the implementation of lethal control within the same site. The intra-site comparison they used,
- also called *before-after control impact* study design, intends to reduce the risks for confounding
- 71 factors of inter-site comparison because these risks are generally higher across space than across
- time. Kutal et al. (2023) was the only study to mix inter-site and intra-site comparisons, by
- combining comparisons before and after hunting season and between sites with and without huntedwelves
- 74 wolves.
- 75 Five of the seven comparative studies on social canids showed that lethal control significantly
- 76 decreased damage recurrence, whereas Santiago-Avila et al. (2018) and Kutal et al. (2023) found no
- 77 significant effect of lethal control on damages. Yet, results suggested variability in the efficiency of
- 78 lethal control. For example, Harper et al. (2008) detected that only the death of an adult wolf male
- 79 was significantly efficient, whereas Bradley et al. (2015) did not detect sex or age conditions on
- 80 efficiency. Variability could also result from the study designs themselves, as these studies used
- 81 different spatial and temporal scales to evaluate lethal control efficiency. Thus, small scale (*e.g.* 4 km
- in Harper et al. (2008)), medium scale (*e.g.* pack territory in Bradley et al. (2015)) or large scale (*e.g.*
- 83 711 km² in Kutal et al. (2023)) could be used for space, sometimes within the same study (*e.g.*
- 84 Santiago-Avila et al., 2018). Time scale could also be disparate, going from several days or months

- 85 (e.g. Blejwas et al. (2002)) to two years (e.g. Kutal et al., 2023) or more (e.g. Bjorge and Gunson
- 86 (1985)). Most importantly, none of the mentioned retrospective studies have both used continuous
- 87 spatial and temporal scales, making it difficult to identify potential gradients of lethal control effects,
- as those suggested in Santiago-Avila et al. (2018). Finally, most of these studies have not controlled
- 89 for livestock presence or abundance, yet the primary factor structuring damage risk and one of the
- 90 highest risks of confounding factor.
- 91 The aim of our study was to evaluate the effects of the lethal control applied for gray wolves in
- 92 France on the intensity of attacks on sheep. We used the data collected by the French
- administrations for the period 2011-2020. Given the retrospective nature of our study, we
- 94 attempted to fill the gaps identified in the previously mentioned retrospective studies on the
- 95 subject. For that purpose, we made intra-site comparisons, used continuous spatial and temporal
- 96 study scales and controlled for livestock presence through an original framework of kernel97 estimation.

98 **2. Methods**

99 2.1 Management context and study area

After historical campaigns of destruction, gray wolves disappeared from France at the beginning of 100 101 the twentieth century, despite being present on the whole metropolitan territory at the end of the 102 eighteenth century (de Beaufort, 1987). From the 1970s, the combination of protective legislation 103 across Europe and the increase of forested habitat allowed the recovery of the species in France 104 through the dispersal of a few individuals from the Apennines population in Italy into the French 105 Alps (Linnell and Cretois, 2018; Fabbri et al., 2007; Poulle et al., 1997). Nowadays, the French wolf 106 population is still expanding but remains mainly restricted to the south-east corner of France, with 107 only a few reproductive packs outside of this area (Louvrier et al., 2018).

- 108 South-east France is characterized by a variety of landscapes, including high altitude mountains,
- 109 Piedmont plains, hills and garrigue lands, where mountain and Mediterranean climates converge.
- 110 Livestock farming is generally extensive. Herds of cow are principally present in the north of the
- region, while those of sheep are present in the whole area (Dobremez et al., 2016). The pastoral
- 112 practice of transhumance, consisting in moving all or part of herds from low to high altitude pastures
- during summer and vice versa during winter, is still important. Duration of grazing season increases
- following a north-south gradient, with some herds from low altitudes in the south grazing all year round (Grente et al., 2022). Livestock owners can share pastures, and even gather their herds to
- 116 limit the cost of transport for transhumance and of herding. In the Alps, herds of cows are generally
- 117 not mixed with flocks of sheep, but the latter can be occasionally mixed with goats. Mean size of
- 118 transhumant flocks of sheep varies greatly between alpine departments and ranges between less
- than 250 sheep up to 2000 sheep (Agreste, 2021).
- 120 To prevent attacks, France subsidies non-lethal measures for protecting flocks of sheep or goat
- 121 through livestock guarding dogs, fences and additional shepherds. Lethal removals by ground
- shooting are authorized in France through derogations to the strict protection provided by the Bern
- 123 Convention and of the Habitats Directive to *Canis lupus*. Other forms of removals (*e.g.* trapping,
- aerial shooting) are prohibited. Conditions of application of ground shooting are defined by
- 125 ministerial decrees (e.g. Ministère de la Transition Ecologique et Solidaire, 2018) and depend on the
- 126 implementation of non-lethal protective measures and on the recurrence of attacks. If the
- 127 conditions are met, the administration assigns the derogation to the livestock owner for a duration
- 128 between two months and five years. The owner itself or a mandated hunter can shoot wolves during

129 an attack on its livestock (hereafter 'simple defense'). When attacks persist, governmental agents or

130 special volunteers who are trained to shoot wolves may support owners in protecting their livestock

131 (hereafter 'reinforced defense'). If the level of attacks escalates in an area, the administration can

- eventually decide to shoot a specific number of wolves under game hunting conditions (hereafter
- 133 'hunting').
- 134 Within the 2011-2020 studied period, lethal removals were for a large majority executed in south-
- east France. Only three lethal removals were outside this area and concerned lone wolves which
- 136 dispersed far and whose territories were very isolated. Their deaths could only lead to an absence of
- 137 attacks. This context was totally different from the one in the Alpine range where the wolf
- 138 population was settled. Therefore, we restricted our study area to the two south-east regions,
- 139 Provence-Alpes-Côte-d'Azur in its entirety and Auvergne-Rhône-Alpes restricted to its five most
- 140 south-eastern departments, Ardèche, Drôme, Isère, Savoie and Haute-Savoie (Figure 1).



Removal type Simple defense Reinforced defense Hunting Interacting

141

Figure 1. Study area and location within France. Purple points represent locations of the 232 analyzed lethal removals of
 wolves, colored according to the three types of shooting implementation (simple defense, reinforced defense, hunting) and

144 of the 32 groups of 'interacting' lethal removals of wolves that were combined for analysis (darkest color). Pastoral surfaces

145 with or without attacks are shown in green and grey, respectively, after correction for inconsistencies between the pastoral

146 and attack datasets (see part 2.2.4).

Pastoral surfaces Attacked Unattacked

147 **2.2 Datasets**

148 2.2.1 Lethal removals

We used the records of deaths of wolves from 2011-2020 compiled by the regional French
administrations. We used the 362 records of wolves killed by legal shootings and that were
geolocated within the study area. Dates of lethal removals were recorded, but not the hour. The
number of lethal removals increased over time and could be divided into three periods, 2011-2014,
2015-2018 and 2019-2020, that represented respectively 8%, 44% and 48% of the dataset, for
annual means of 7 (± 6), 40 (± 6) and 87 (± 3) lethal removals respectively.

Almost half of the records were cases of reinforced defense (48%), while simple defense and hunting cases represented respectively 27% and 25% of the dataset. There were 58 records (16%) for which no wolf body was found after the shooting, but they were considered as killed because there was enough evidence for lethal injuries. For the remaining 304 records, information from the necropsy reports was available on the sex and the estimated age class (pup, subadult or adult) of the killed wolves and on the annual reproductive status for females.

161 2.2.2 Attacks on sheep

162 We used the claims of attacks on livestock from 2010-2020 compiled by the French administration

and for which wolf responsibility could not be discarded. Among the 25 948 available claims, we

used the 23 645 claims that were geolocated within the study area. We then restricted the dataset

to the 21 936 claims on sheep. Each claim corresponded to one or several dead or wounded sheep.

166 Apart very few specific cases, all claims were checked in the field by trained public agents. They

167 followed a standardized procedure to estimate the date of the attack and if wolf responsibility could

168 be discarded or not. Compensation was attributed regardless of non-lethal measures up to 2019.

From 2020, compensations after two compensated attacks depended on the implementation of
 protective measures, with conditions of implementation varying locally according to the level of

171 attacks of municipalities. Protective measures are widely implemented in south-east France (Meuret

et al., 2020). We have therefore confidence in the level of reports of attacks noticed by breeders.

173 2.2.3 Spatio-temporal sheep distribution

174 We used the data from the pastoral census of the two south-east regions Provence-Alpes-Côted'Azur and Auvergne-Rhône-Alpes conducted by the National Research Institute for Agriculture, 175 176 Food and the Environment (Dobremez et al., 2016). The census indexed and delimited for the period 177 2012-2014 the pastoral units at medium or high altitude used by flocks, as well as the pastoral zones 178 at lower altitudes. Giving the relative stability of the use of the pastoral units and zones over time 179 (Grente et al., 2022), we assumed the information of the pastoral census was valid for our entire 180 study period. Pastoral units and zones delimited the spatio-temporal area where attacks could have 181 occurred. Therefore, our analysis was restricted to the pastoral area delimited by this census, in 182 order to control for livestock presence.

183

184 Among the 3576 pastoral units and the 8862 pastoral zones of the census, we used the 3176 units

and 8184 zones within the study area. We restricted the dataset to units and zones that were used

186 for grazing sheep, regardless of the other livestock. The annual number of grazing sheep was on

187 average 955 (standard deviation, hereafter ' \pm ', 764) for units (range 2–6000) and 412 (\pm 411) for

zones (range 1—3600). The months for which the area was grazed were known for each pastoral
 unit. Only the seasons in use were available for pastoral zones. Therefore, we attributed to the

190 pastoral zones the months corresponding to the seasons for which they were grazed (spring: March

to May, summer: June to August, fall: September to November, winter: December to February). The
annual number of grazed months was on average 6.9 (range 1—12). The total size of our final
dataset of pastoral units and zones, hereafter pastoral units, was 4987 (Figure 1).

194

195 2.2.4 Attributing attacks to pastoral units

The attack dataset showed 59% of consistency with the dataset of pastoral surfaces, meaning that 59% attacks were within pastures that were in use according to the pastoral census. The other attacks could be outside pastoral surfaces, or could have occurred in months for which pastoral surfaces were not supposed to be in use according to the pastoral census. These mismatches could result from changes in the spatial or temporal use of pastoral surfaces, from errors during the attack record (*e.g.* geolocation approximation), or because attacks occurred in non-indexed pastoral surfaces.

- 203 When attacks were within pastures and when difference between the month of their dates and the
- sets of grazed months of their pastures did not exceed three months (9% of total attack dataset), we
- 205 only updated the pastoral census by adding the months of the attacks to the corresponding
- 206 pastures. Otherwise, we relocated attacks to their closest pastures within a 500 meters radius by
- 207 favoring pastures with the smallest difference between their sets of grazed months and the months
- 208 of the attacks. When these differences exceeded 3 months, or if no pastoral surface could be found
- within a 500 meters radius around the attacks, we excluded these attacks (13% of total attack
- dataset). Relocations were done at the closest point of the pastoral surfaces to the original attack
 locations (19% of total attack dataset). When necessary, we added the months of the relocated
- 212 attacks to the sets of grazed months of their new pastoral surfaces. The final dataset of attacks had
- 213 19 302 events.
- Contrary to attacks, lethal removals did not require to be located within pastoral surfaces, as huntingdid not necessarily occur within pastoral surfaces.

216 **2.3 Analysis**

Apart from being precisely geolocated, our data about attacks and lethal removals were very
unbalanced in favor of attacks (19 302 attacks against 362 lethal removals). Consequently, we could
adopt an original approach consisting in choosing removals as focal points through the use of 3D
buffers, and in computing attack intensities over continuous spatial and temporal scales.

221 2.3.1 Construction of 3D buffers

We attributed to each lethal removal a spatio-temporal buffer, which can be visualized as a 3D cylinder centered on the lethal removal event. The circle base of the cylinder corresponded to the spatial dimensions (X, Y) and had the role to scan for attacks within the pastoral space around the lethal removal. The height of the cylinder corresponded to the temporal dimension, the left and right parts scanning for attacks that had occurred before or after the lethal removal, respectively (Figure 2.A).



229 230

Figure 2. Data transformations for one lethal removal (purple point) occurring on 15th of August 2017 as an example. A. 231 Attribution of the 3D spatio-temporal buffer around the removal. Each slice represents the occurrence of observed attacks 232 (red points) in a radius of 10 km around the removal in a given month m. Months could be complete (from m_{-2} to m_2) or 233 partial (m₋₃ and m₃, here only 15 days each). Monthly slices were chosen for visual convenience, but slices were daily in the 234 analysis. Attacks (red points) in the buffer have 3 coordinates: 1 temporal (height of the cylinder), and 2 spatial (x,y). B. 235 Reduction of 2D spatial coordinates of observed attacks of the 3D buffer to 1D, through the Euclidian distance to the 236 removal (coordinate i, i.e. rows). Time coordinate becomes coordinate j (columns). C. Dividing one 10-km radius slice around 237 the removal into 200 rings (here only 50 are represented for visual convenience) to calculate the total pastoral area pa that 238 was used during the corresponding month of the slice. The column of pa is then duplicated into the number of days of the 239 corresponding month (i.e. information about pastoral use was monthly). The operation is repeated for each slice of the 3D 240 buffer. D. Simulation of coordinates i and j of attacks (green points) based on the matrix of pa from step C. The number of 241 simulated attacks is the same as the number of observed attacks. The size of points represents the spatial correction $1/pa_r$ 242 based on similar calculations as step C but achieved on the pastoral surfaces regardless of their use (i.e. as in the most left 243 circle). Note that the observed and simulated datasets are placed on the same figure for visual convenience here, but were 244 treated separately during the kernel estimations.

245 The spatio-temporal buffer was composed solely of the pastoral surfaces where attacks could occur.

246 Therefore, the buffer was sparse, and remained empty where there was no pastoral surface around 247 the removal event, or when the pastoral surface was not grazed according to the pastoral census.

248

249 We set the radius of the cylinder at 10 km, in order to encompass more than one wolf territory

250 radius, estimated at 7–8 km in France through telemetry and genetic tracking (Duchamp et al.,

251 2012). Therefore, ripple effects, if any, could be observed up to 10 km from the lethal removal. The 252 pastoral surfaces overlapping the spatial buffer were truncated to the spatial buffer.

253

254 We set the height of the cylinder to 180 days, meaning that all attacks occurring 90 days before and 255 90 days after the lethal removal were part of the temporal buffer. Thus, we analyzed the effect of 256

lethal removal on a full season if the removal event occurred in the middle of the season, or on two

257 successive seasons if not. The day of the removal was not part of the buffer, because we could not know if an attack recorded at the day of the removal occurred after of before the removal, having no 258

- 259 information about the hour of both attacks and removals.
- 260

261 Attacks were not homogeneously distributed across space and time (Grente et al., 2022). Increasing

262 the radius or the height of buffers compressed the attack distribution towards the centers of buffers 263 (*i.e.* at small spatial and temporal distances), and prevented an optimal reading of the results.

264

265 We ensured that the buffers did not overlap geographic areas or periods without data about attacks.

266 We removed the 84 lethal removals occurring at less than 10 km from the terrestrial limits of the study area or occurring after the 1st October 2020, eventually leaving 278 analyzed lethal removals
(Figure 1).

- 269 Most buffers overlapped in space and in time, with a total of 86% of buffers overlapping with at least
- 270 another one. However, only few removal events were simultaneously close in space and time (Figure
- 271 3), with 82 removals separated by less than 5 km and less than 25 days. In these cases, we
- 272 considered that the risks of interaction between removals were high enough. We called these
- 273 removals, 'interacting removals'. We pooled the interacting removals into 36 groups, with each
- 274 corresponding to a set of interacting removals. For the special case where two removals were not
- 275 interacting, but were both interacting with a third removal, we considered these three removals
- were part of the same group of interacting removals. On average, 2.28 wolves (\pm 0.45, range 2–3)
- 277 were killed per group.
- 278 We attributed a single buffer to each group of interacting removals, of 10 km radius, of 180 days
- 279 height and centered on the centroid of the interacting removals of the group. To compare the attack
- intensity before and after each group of interacting removals, we excluded the attacks recorded
- between the date of the first removal of the group and the date of the latest removal of the group.
- 282 The exclusion period lasted 7.5 days on average (± 8.38, range 0–27). In cases where interacting
- removals occurred the same day, only the date of the removals was not part of the buffer, as
- 284 previously.
- 285 We ended up with 232 spatio-temporal buffers, including 36 buffers of groups of interacting
- removals, corresponding to a total of 6110 attacks. On average, a buffer included 26.3 attacks (±
- 287 20.7, range 0—100).
- 288



Figure 3. Spatial and temporal distances between the 352 pairs of lethal removals whose spatio-temporal buffers
 overlapped.

291 2.3.2 Reducing the number of spatial dimensions

292 Methods for the analysis of interactions of several types of datasets with three dimensions are yet to

293 be developed and may suffer from the curse of dimensionality (Wikle et al., 2019). Therefore, we

- followed standard geostatistical practices and reduced the two spatial dimensions of the buffers to
- 295 one by computing the Euclidian distance of each attack to the center of the buffer. In other words,

296 we converted the spatial circle of buffers to a line. The temporal dimension corresponded to the 297 difference in days between each attack and the date of the removal of the buffer. In the case of

298 interacting removals, we used the first (or last) date of the group if the attack occurred before (or

after) the exclusion period. Thus, each attack occurring in a buffer had then two coordinates (i, j)

300 corresponding to the spatial *i* and temporal *j* distances from the center of the buffer (Figure 2.B).

Nevertheless, converting a 2D circle into a 1D line always induces a spatial bias that gives more

302 representation of points farther away, simply because the circle perimeter increases linearly with

303 circle radius. Therefore, we should have applied a spatial correction of 1/*i* to each attack of a buffer,

with *i* the Euclidian distance of the attack to the buffer center. However, the buffers were delimited

305 by pastoral surfaces that were not homogeneously distributed around the buffer center. In other

306 words, the spatial bias did not increase linearly with the radius, but according to the perimeter 307 fraction that intersects with pastoral surfaces. For example, the spatial bias could be low at the

fraction that intersects with pastoral surfaces. For example, the spatial bias could be low at the
 furthest distances of the buffers if pastoral surfaces were almost non-existent at these distances.

309 To incorporate pastoral surfaces to the spatial correction, we sliced up the circle of each buffer into

310 200 rings *r* of progressive radius, from 0.025 to 9.975 km. Width of rings were therefore 50 m. We

calculated *pa*_{*r*}, the area of pastoral surfaces of each ring *r*. We then attributed to each attack from

312 the buffer the inverse of the pastoral area of its ring, $1/pa_r$ (Figure 2.B-D), corresponding to its

weight in the later analysis (part 2.3.4). Because attacks could only be located in pastoral surfaces

314 (part 2.2.4), *pa_r* of rings containing attacks could not be null.

315 2.3.3 Simulating attacks according to livestock presence

316 We controlled for livestock presence in each buffer by randomly simulating attacks with a

317 distribution defined according to the locations of pastoral surfaces of the buffer and their grazed

318 months. The simulated dataset corresponded to the expected distribution of attacks if they were

319 only dependent on the livestock presence.

For each buffer, we used the same 200 rings *r* as in part 2.3.2. For each month *m* represented in the temporal buffer, we calculated the set of *pa_{c,m}* by considering only pastoral surfaces that were

grazed this month according to the pastoral census (Figure 2.C). Thus, we constructed a pa matrix [r, r]

m] for each buffer, where columns were sorted following the order of months in the temporal

buffer. Each column *m* was then duplicated according to the number of days of this month included

in the temporal buffer. Thus, the matrix dimensions were 200 rows and 180 columns. Eventually, the

326 matrix coordinates (r, m) became similar to the coordinates (i, j) of the observed attacks, with a

327 slightly coarser spatial resolution of 50 meters (*i.e.* ring width) and the same temporal resolution of 1

328 day.

329 We converted the matrix into a two-dimensional pixel image. We generated a random point pattern,

330 with the pixel image acting as the probability density of the points. We generated the same number

of points as the number of observed attacks in the buffer. We extracted the coordinates (i, j) of each

point, hereafter called simulated attack, corresponding to its spatial and temporal distances to the

buffer center. We also attributed the spatial correction $1/pa_r$ for each simulated attack, following

the same calculations as in 2.3.2 (Figure 2.D). We repeated the simulation 1000 times for eachbuffer.

336 2.3.4 Kernel estimates

For a subset of buffers (see subset definition in the next section 2.3.5), we applied an anisotropic Gaussian kernel estimation of the observed attacks, based on the distribution of their spatial *i* and

- temporal *j* coordinates, with the spatial correction $1/pa_r$ as weight. We used two separate
- 340 bandwidth parameters for spatial distance and temporal lag, of 500 meters and 2 days. The choice of
- 341 the bandwidth parameters was *ad-hoc*, and met the needs of adopting a resolution close to the
- 342 matrices used to simulate attacks (i.e. cells of 50 meters and 1 day) while smoothing attack
- intensities enough in order to obtain interpretable results. We also considered the spatio-temporal
- distributions of attacks before removals ($j \in [1, 90]$ and after removals ($j \in [91, 180]$) as two separate
- 345 distributions by appropriately right- and left-truncating kernels in the time dimension for attacks
- occurring before and after the date of removals, respectively. We corrected kernel estimates for
- outliers by capping the estimated observed attack intensities $\lambda_{obs}(i, j)$ to $Q_3 + 1.5 \times (Q_3 Q_1)$, where Q_1
- and Q_3 were the first and third quantiles of the estimated distribution of λ_{obs} (Walfish, 2006).
- 349 We repeated the Gaussian kernel estimation for the simulated buffers in the same conditions. To
- this end, we combined the simulated attacks of the same subset of buffers for each of the 1000
- 351 simulations. The kernel estimation was applied on each of the 1000 combined simulations. We then
- calculated the mean estimated simulated attack intensities $\lambda_{mean_sim}(i, j)$ of the 1000 kernel
- 353 estimations. The mean kernel estimation obtained corresponded to the expected attack intensities
- 354 given pastoral use.
- 355 To control the observed kernel estimation for livestock presence, we calculated the corrected
- 356 intensities $\lambda_{corr}(i, j)$ as:

$$\lambda_{corr}(i,j) = \lambda_{mean_{sim}}(i,j) - \lambda_{obs}(i,j)$$
(1)

358 Thus, a value of $\lambda_{corr}(i, j)$ lower, equal or higher than zero indicated that the observed attack intensity

- at coordinates (*i*, *j*) was lower, equal or higher, respectively, than the attack intensity expected according to pastoral use.
- We calculated the general trends of observed or corrected attack intensities at different spatiotemporal scales *s* as:

363
$$Y_{\lambda,s} = \frac{\sum_{i=1,j=91}^{i=I,j=91+J} \lambda'(i,j) - \sum_{i=1,j=90-J}^{i=I,j=90} \lambda'(i,j)}{(I \times J)/2} \times 100$$
(2)

- 364 where I and J were the number of rows (spatial distance to removals) and columns (number of days 365 before or after removals) of scale s, $\lambda'(i, j)$ was the observed or corrected intensity at the coordinates 366 (*i*, *j*) rescaled over [0,1] for observed intensity or over [-1,1] for corrected intensity by using the 367 minimum and maximum values of $\lambda_{obs}(i, j)$ or $\lambda_{corr}(i, j)$ over the temporal range 90-J, 91+J. The 368 rescaling allowed trend comparison between subsets of buffers. The division by $(I \times J)/2$ allowed trend comparison between scales s of different sizes. We investigated trends of three spatio-369 370 temporal scales s: the small scale (s = 1) of spatial range 0-2 km and of temporal range ± 15 days (I = 371 40, J = 14), the intermediate scale (s = 2) of spatial range 0-6 km and of temporal range ± 45 days (I = 372 120, J = 44) and the large scale (s = 3) of spatial range 0-10 km and of temporal range \pm 90 days (I = 373 200, J = 89).
- 374 In addition, we investigated the potential spatial shift of attack intensities at large scale s = 3 by
- applying Equation 2 separately for each spatial coordinate *i* ($i \in [0,200]$). There was no need to
- adjust results by $(I \times J)/2$ as we only investigated shift at s = 3. Namely, we transformed Equation 2
- 377 into Equation 3: $Y_{\lambda}(i) = (\sum_{j=91}^{180} \lambda'(i,j) \sum_{j=1}^{90} \lambda'(i,j)) \times 100.$
- 378

380 2.3.4 Investigated categories of removals

381 The subsets of buffers on which we applied the kernel calculations were divided into 2 general

- 382 categories. The first category corresponded to the whole subset of interacting groups of removals
- 383 ('*Interacting*', n = 36, corresponding to 82 lethal removals). The second category corresponded to
- the whole subset of removals that were not interacting ('Non-interacting', n = 196).



Figure 4. Topography of the study area, with the 8 highlighted zones from the Zone subcategory that were selected for the analysis. Upward and downward triangles represent lethal removals that could be attributed or not to one of the 27 zones, respectively. Colors show the subset from the Season subcategory.

- 385 We also divided the second category into 4 subcategories. The first subcategory corresponded to
- subsets defined according to removal locations. We delimited zones by using 25 Alpine mountain
- 387 ranges defined by hydrography and orography (Knopf, 2019) and the limits from OpenStreetMap of
- 388 the protected area *Parc Naturel Régional des Préalpes d'Azur* (PNRPA) and of the military zone
- 389 *Canjuers*, both in the south of the study area. We attributed removals to their closest zone to a limit
- of 15 km. We selected the 8 subsets with at least 10 removals (Figure 4): PNRPA (n = 33),
- 391 Mercantour (23), Vercors (18), Les Trois-Evêchés (18), Vanoise (15), Préalpes de Digne (12), Ecrins
- 392 (12) and *Canjuers* (11).
- 393 The second subcategory corresponded to the 5 subsets defined according to the wolf seasons during
- which removals occurred (Figure 4): *Winter* (November December, n = 30), *Reproduction*
- 395 (February March, 9), *Denning* (April June, 35), *Rendez-vous* (July August, 59) and *Fall*
- 396 (September October, 63).
- 397 The third subcategory corresponded to the 3 subsets defined according to removals administrative
- types (Figure 1): *Simple defense* (n = 69), *Reinforced defense* (75) and *Hunting* (52).

- 399 The fourth subcategory corresponded to the 2 subsets defined according to the sex of the killed wolf
- 400 when known: *Male* (n = 83) and *Female*, (n = 78). At the necropsy, 7 females of our final dataset of
- 401 lethal removals were declared as being reproductive at least the year they were killed, which we
- 402 considered too low to constitute a subset. Annual reproductive status of males was unknown.
- In addition, we considered that the distribution of lethal removals among the three age classes when
 known was too unbalanced to investigate a potential effect of age class (177 adults, 24 subadults
 and 16 pups).

406 **2.4 Implementation**

- 407 The analysis was conducted in R 4.2.3 (R Core Team, 2023). We used the R packages *tidyverse 2.0.0*
- 408 (Wickham et al., 2019), readxl 1.4.2 (Wickham and Bryan, 2023), and sf 1.0-12 (Pebesma and Bivand,
- 409 2023) to handle the datasets. We used the R package *spatstat 3.0-6* (Baddeley et al., 2015) to
- simulate point patterns in part 2.3.3. We adapted the function *kde2d.weighted* from *ggtern 3.4.2*
- 411 (Hamilton and Ferry, 2018) to separate the kernel estimates before and after the removal date in
- 412 part 2.3.4. We used the package *osmdata 0.2.3* (Padgham et al., 2017) to extract limits of PNRPA and
- 413 of Canjuers military zone. We used the packages *RColorBrewer 1.1-3* (Neuwirth, 2022), *tidyterra*
- 414 0.4.0 (Hernangómez, 2023), geodata 0.5-8 (Hijmans et al., 2023), scales 1.2.1 (Wickham and Seidel,
- 415 2022), *ggfx 1.0.1* (Pedersen, 2022a), *ggnewscale 0.4.9* (Campitelli, 2023), *metR 0.14.1* (Campitelli,
- 416 2021), patchwork 1.1.2 (Pedersen, 2022b) and ggtext 0.1.2 (Wilke and Wiernik, 2022) for data
- 417 vizualisation. The datasets, codes and results are available online (Grente et al., 2024).

418 **3. Results**

- 419 Trends in the observed and corrected attack intensities for all analyzed subsets and for all spatio-
- 420 temporal scales *s* are displayed in Table 1. Distributions of these trends are displayed in Figure 5.
- 421 Heatmaps of observed and corrected kernel estimates of subsets are displayed in Figure 6 and Figure
- 422 S1. Trends as a function of spatial coordinate *i* are displayed in Figure 7 for all subsets.
- 423 Subsets generally presented a general pattern of observed or corrected intensities that formed
- 424 patches of high intensities at low distances from removals, below 2 km (e.g. Non-interacting [Figure
- 425 6.A], Denning [Figure 6.F], Reinforced defense [Figure 6.J]). Some subsets also showed patterns that
- 426 progressively moved away from this general pattern, starting from subsets showing supplementary
- 427 patches of high intensities at higher spatial distances such as *Fall* (Figure S1.F), to subsets splitting up
- 428 their patches of low and high intensities at all analyzed scales such as *Ecrins* (Figure S1.B).
- 429 Trends in totals of attack intensities leaned towards a decrease of intensities after removals, with
- 430 their main modes located between -21% and -9% (Figure 5). However, this general decrease
- 431 concealed disparities between subsets. Indeed, from subset trends in Table 1 or from heatmaps, we
- 432 could observe three main possible types of attack intensity changes after lethal removals, *i.e.*
- 433 decrease, no change and increase of attack intensities. Most importantly, the results did not only
- 434 vary according to subsets, but also according to the analyzed scales *s*, to the correction of analysis
- 435 for livestock presence and to the spatial distance to lethal removals.
- 436 3.1 At small scale

- 437 At small scale *s* = 1, 17 subsets over 20 showed an observed negative trend of attack intensity after
- 438 removals (Table 1). Only the subsets of *Canjuers* and *Hunting* showed a positive observed trend after
- 439 removals (Table 1). *PNRPA* was the only subset showing an almost zero observed trend.
- 440 The control for livestock increased the strength of two-thirds of negative trends and of all positive
- 441 trends at scale 1 (Table 1). Thus, the corrected trend distribution flattened the observed trend
- 442 distribution at scale 1, and reinforced its distribution tails (Figure 5).

443 **3.2 At intermediate scale**

- 444 Results at the intermediate scale s = 2 showed the same proportions of observed negative, positive 445 or zero trends as those at scale 1. However, the classification of subsets within each type of result
- 446 could change. For example, the subset of *Hunting*, formerly showing a weak but positive trend (4%)
- 447 at scale 1, was the second subset showing the highest negative trend (-40%) at scale 2. Conversely,
- 448 Vercors, formerly the fifth subset showing the highest negative trend (-35%) at scale 1, had the
- second highest positive trend (11%) at scale 2. *Ecrins* was now the zero-trend subset. Others were
- 450 consistent across scales, such as *Préalpes de Digne* that remained in the first three subsets showing
- 451 the highest negative trends at both scales.
- 452 Contrary to scale 1, the control for livestock presence at scale 2 reduced the strength of half of
- 453 negative and of two-thirds of positive observed trends (Table 1). Thus, the main mode of the
- 454 corrected trend distribution at scale 2 increased from -21% to -9%, and its right tail extended toward
- 455 higher values than the observed trend distribution (Figure 5).

456 3.3 At large scale

- 457 At large scale s = 3, the subsets showing an observed negative trend were still dominant, with 15
 458 subsets concerned over 20, whereas three subsets showed an observed positive trend. The observed
 459 zero trend subsets were again different from the other scales: *Vercors* and *Simple defense* (Table 1).
- 460 The differences between the observed and corrected intensities were important at scale 3. Thus, the
- 461 control for livestock presence changed the trend sign in 5 subsets over the 20 analyzed, for only 2 at
- scale 1 and 1 at scale 2 (Table 1). The control for livestock at scale 3 reduced the strength of two-
- 463 thirds of negative and of positive trends (Table 1).
- 464 Thus, the main mode of the trend distribution at scale 3 increased from -15% to -9.5% when
- 465 corrected, but contrary to scale 2, the main mode narrowed instead of flattening (Figure 5). In
- 466 addition, two minor modes emerged when using corrected intensities, at -31% and at 20%. In other
- 467 words, control for livestock exacerbated the emergence of three types of results at scale 3: a
- 468 moderate decrease, a small decrease and a moderate increase. The subset of *Reproduction*
- 469 corresponded to the negative mode of -31%, and the subsets of Vercors, Préalpes de Digne and
- 470 *Winter* particularly contributed to the positive mode of 20% (Table 1).

471 **3.4 Differences between observed and controlled intensities in heatmaps**

- 472 Some subsets showed large differences between the panel ii of observed intensities and the panel iv
- 473 of corrected intensities (Figures 6 and S1), such as *Préalpes de Digne* (Figure 6.D), *Vercors* (Figure
- 474 6.E), *Reinforced defense* (Figure 6.J), *Winter* (Figure 6.H) or *Female* (Figure S1.I).
- 475 Yet, not all of these subsets were associated with high differences between observed and corrected
- 476 trends at scale 3. For example, the subset of *Reinforced defense* had a trend difference of 5%,
- 477 against 54% for the Winter subset. In the case of Reinforced defense subset, the effect of livestock
- 478 control was the same before and after removals (*i.e.* time 0, Figure 6.J), whereas in the case of

479 *Winter* subset, the correction decreased the intensities before removals and increased those after

480 removals (Figure 6.H). These differences lay on the monthly distribution of removals. Corrections

varied more greatly along the temporal axis when removals occurred at the limit between an intense

and calm grazing period, like removals of the *Winter* subset (Figure 6.H.ii).

483 **3.5 Attack intensities by spatial distances**

484 When looking at the observed attack intensity trends as a function of spatial distance to removals,

three types of results were observed, with balanced proportions (Figure 7): A decrease or an

486 increase over all distances (6/20 subsets, e.g. Interacting subset), phases of decrease interrupted by

487 phases of weak increase (7/20 subsets, *e.g. Male* subset) and phases of decrease interrupted by

488 phases of strong increase (7/20 subsets, *e.g. Ecrins* subset).

489 Repeating the same exercise for corrected attack intensity trends led to the same three types of

490 results, but with a large majority of subsets presenting decrease phases interrupted by strong

491 increase phases (14/20 subsets) and only two subsets, Vercors and Winter, presenting no trend sign

492 shifts over all spatial distances.

493 Most of the time, trends shifted from negative to positive values, sometimes several times. *Vanoise*

494 and *Male* were the only subsets to show an observed or corrected positive trend after removals at 0

495 km while presenting trend shifts. Distances at which first trend shifts appeared ranged from less

496 than 1 km (*Ecrins*) up to 7 km (*Hunting*). On average, 1.7 trend shifts occurred over 10 km.

497 The two subsets of *Ecrins* and *Canjuers* showed a multimodal distribution of trend shifts, with shifts 498 equally apart of approximately 1 km. In these cases, as in most subsets presenting trend shifts, there

499 was no symmetry between negative and positive phases, *e.g.* there could be an intense decrease in

500 attack intensity after removals at low distances, and only one or several weak increases in attack

- 501 intensity after removals at higher distances, such as the subsets of *Vanoise* and *Reproduction*.
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516 Table 1. Trends (%) of observed attack intensities $Y_{\lambda_{obs},s}$ and of corrected attack intensities $Y_{\lambda_{corr},s}$ after lethal removals

517 calculated over three different spatio-temporal scales s, 1 (0-2 km, \pm 15 days), 2 (0-6 km, \pm 45 days) and 3 (0-10 km, \pm 90

518 days) following Equation 2, separately for each subset. All trends are comparable because intensities were rescaled and

trends were relative to the number of involved rows and columns. Column "Total" indicates the number of lethal removals
of each subset or of groups of lethal removals for the category "Interacting". Column "Figure" gives the corresponding

521 figure number of heatmaps. Gradient of colors ranges from dark blue for the lowest value to dark red for the highest value,

522 while the middle color in white corresponds to zero.

| | | | | | Sca | ale 1 | Sca | ale 2 | Sca | ale 3 |
|-------------|--------------------|-----------------------|-------|--------|-----------------------|------------------------|-----------------------|------------------------|-----------------------|------------------------|
| Category | Subcategory | Subset | Total | Figure | $Y_{\lambda_{obs},1}$ | $Y_{\lambda_{corr},1}$ | $Y_{\lambda_{obs},2}$ | $Y_{\lambda_{corr},2}$ | $Y_{\lambda_{obs},3}$ | $Y_{\lambda_{corr},3}$ |
| | | | 196 | 6.A | -16 | -14 | -28 | -23 | -18 | -8 |
| | Geographic zone | PNRPA | 33 | 6.C | 1 | -1 | -9 | -15 | -10 | -15 |
| | | Mercantour | 23 | S1.D | -39 | -54 | -27 | -23 | -21 | -10 |
| | | Trois Evêchés | 18 | S1.C | -34 | -29 | -24 | -3 | -19 | 6 |
| | | Vercors | 18 | 6.E | -35 | -38 | 11 | 31 | 0 | 19 |
| | | Vanoise | 15 | S1.E | -18 | -30 | -27 | -38 | -5 | -11 |
| | | Ecrins | 12 | S1.B | -15 | -21 | 0 | 2 | 3 | 9 |
| | | Préalpes de Digne | 12 | 6.D | -77 | -67 | -30 | -2 | -23 | 21 |
| | | Canjuers | 11 | S1.A | 6 | 13 | 2 | 0 | -3 | -8 |
| Non- | Season | Fall | 63 | S1.F | -17 | -17 | -48 | -49 | -33 | -11 |
| interacting | | Rendez-vous | 59 | 6.G | -10 | -16 | 12 | 9 | 7 | -8 |
| | | Denning | 35 | 6.F | -62 | -96 | -26 | -42 | 2 | -12 |
| | | Winter | 30 | 6.H | -33 | -17 | -16 | 25 | -23 | 31 |
| | | Reproduction | 9 | S1.G | -23 | -40 | -8 | -15 | -17 | -31 |
| | Class | Reinforced defense | 75 | 6.J | -14 | -15 | -8 | -3 | -11 | -6 |
| | | Simple defense | 69 | 6.1 | -20 | -27 | -19 | -26 | 0 | -3 |
| | | Hunting | 52 | 6.K | 4 | 20 | -40 | -39 | -33 | -16 |
| | Sex | Male | 83 | S1.H | -21 | -25 | -19 | -12 | -14 | -3 |
| | | Female | 78 | S1.I | -5 | 1 | -15 | -10 | -13 | -2 |
| Interacting | | | 36 | 6.B | -44 | -88 | -14 | -28 | -7 | -13 |

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Figure 5. Density distributions of the trends of observed (dotted lines) and corrected (solid lines) attack intensities after lethal removals at the spatio-temporal scales 1 (0-2 km, \pm 15 days), 2 (0-6 km, \pm 45 days) and 3 (0-10 km, \pm 90 days) presented in Table 1 (n = 20 subsets per distribution).



A. Non-interacting lethal removals







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541 Figure 6. Results of kernel estimation of attack intensities surrounding lethal removals, detailed for each analyzed subset of 542 removals (all panels except B.) or groups of removals for the subset of interacting removals (B.). Each panel includes (i) the 543 map of the removal locations or of the centroids of the groups, with the annotated number of involved removals or groups, 544 (ii) the monthly distribution of the removal dates or of the middle dates of the groups, (iii) the observed attack intensities 545 λ_{obs} and (iv) the corrected attack intensities λ_{corr} . Intensities are given along the continuous temporal axis (x-axis) and spatial 546 axis (y-axis). The coordinate (0,0) indicates the date and the location at which removals or groups of interacting removals 547 occurred. Grey rectangles of subpanels iii and iv delimit the three analysis scales 1 (0-2 km, ± 15 days), 2 (0-6 km, ± 45 days) 548 and 3 (0-10 km, \pm 90 days). A value of λ_{corr} lower, equal or higher than zero indicated that the observed attack intensity was

lower, equal or higher, respectively, than the attack intensity expected according pastoral use. (Part 1/4)





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553 Figure 6. (Part 3/4)





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561 **4. Discussion**

We evaluated the effects of lethal control of wolves on attack intensity distribution through a kernel
 density estimation that had never been used before for this topic. Our method contrasted with the
 Cox proportional hazard models commonly used by previous studies to estimate the risk of

recurrence of a single attack over a fixed area (Bradley et al., 2015; Harper et al., 2008; Santiago-

566 Avila et al., 2018).

567 We conducted the analysis on 20 subsets of lethal removals of wolves in France, for a total of 278

568 analyzed lethal removals occurring between 2011 and 2020. Subsets were based on interactions

569 between removals, administrative type, location or date of removals, or on sex of killed wolves. We

570 corrected attack intensities by sheep presence according to a pastoral census from 2012-2014. We

571 conducted the analysis over 10 km around removals and 90 days before and 90 days after removals.

572 This method provided an efficient visualization of attack intensity distribution in the form of

573 heatmaps over continuous spatial and temporal scales. We also computed attack intensity trends

Figure 7. Trends (%) of observed attack intensities $Y_{\lambda_{obs}}$ (dotted line) and of corrected attack intensities $Y_{\lambda_{corr}}$ (solid lines) after lethal removals calculated along the spatial scale of 10 km with a fixed temporal scale of ± 90 days following Equation 3, separately for each subset. Top facet strips indicate the category or subcategory of the subset, and the bottom facet strips indicate the subset name. All trends are comparable because intensities were rescaled.

after removals and added two smaller spatio-temporal analysis scales (0-6 km and ± 45 days; 0-2 km
and ± 15 days) to the computations.

576 4.1 Common features among subsets and scale influence

Heatmaps revealed a high variability of patterns in spatio-temporal distributions of attacks across
subsets (Figure 6, Figure S1), although two features tended to be common to all subsets. First,
removals were generally preceded by high attack intensities at their locations, meaning they were
applied in response to attacks, in accordance with French legislation. Second, attacks tended to
aggregate below 2 km from removal locations, regardless of time and of other possible aggregation
spots. Therefore, the attack process was generally a persistent phenomenon at these locations,

although showing possible interruptions in time, even before removals (*e.g.* subset of *Rendez-vous*).

- 584 Trends of attack intensities once removals occurred (Table 1) and their distributions (Figure 5) broke
- 585 down the variability observed in heatmaps into three possibilities: Decrease in attack intensity, zero
- 586 or almost zero trend, and increase in attack intensity. Proportions of each type varied according to
- 587 the analysis scale and the use of livestock control, but decrease in attack intensity remained over-
- represented in all cases (75-85% of subsets). Nevertheless, for about half of the subsets, the greater
- the analysis scale, the lesser the decrease. Thus, 96% of decrease in corrected attack intensities were
- 590 observed for the *Denning* subset at the smallest scale (0-2 km, ± 15 days), against only 12% of
- decrease at the largest scale (0-10 km, \pm 90 days). If considering the uncorrected results, this was
- 592 even worse, as the 62% of decrease at scale 1 became a 2% increase at scale 3 for this subset.
- Conversely, the three subsets showing a zero trend or an increase of attack intensities after
 removals at scale 1 showed negative trends at larger scales (*PNRPA, Canjuers, Hunting* subsets). The
- 595 other half of subsets presented a non-linear effect of scales, with intermediate scale 2 (0-6 km, ± 45
- 555 other half of subsets presented a non-inteal effect of scales, with intermediate scale 2 (0-0 km, ± 4
- 596 days) showing the highest or lowest trends (*e.g. Rendez-vous, Non-interacting* subsets). Overall, the 597 effects of lethal control according to spatio-temporal analysis scale were difficult to predict.

598 **4.2 Ripple effects of lethal removals**

599 When focusing on the variation of attack intensities after lethal removals along the spatial scale of 600 10 km with a fixed temporal scale of ± 90 days (Figure 7), we frequently observed localized positive 601 trends, although this distance greatly varied across subsets.

- 602 Shifts in the location of attack could result from pack dissolution (Cassidy et al., 2023), especially if 603 killed wolves were breeders. Yet, only 7 killed females among the 105 necropsied females were 604 evaluated as breeders in our dataset. Nevertheless, this number could have been underestimated, 605 and reproductive status of males were unknown. In addition, the analyzed spatial and temporal 606 scale (10 km, 90 days) may also not be enough considering the large dispersal abilities of wolves 607 (Mech, 2020), although this spatial scale was sufficient for Santiago-Avila et al. (2018) to detect 608 potential attack shifts. In our case, the persistence of high intensities of attacks below 2 km from 609 removals, frequently quickly after removals and on the long term, weakened the pack dissolution 610 hypothesis, as settlement of new wolves was unlikely to be as fast, even in the dense population of 611 the Western Alps.
- 612 Remaining wolves from targeted packs could also have shifted part of their attacks toward other
- 613 flocks, possibly because of removals without necessarily implying pack dissolution. This shift could
- also simply result from changes in the spatio-temporal distribution of sheep flocks. Control for
- 615 livestock presence was designed to correct for this bias. However, correction was applied on a
- 616 monthly basis. Shepherds may change the grazing zone or pastoral surface of their flocks from one
- 617 week to another. These changes would not be corrected by our monthly correction for livestock

presence, especially since our correction was based on a ten-yearly pastoral census and may not be
up-to-date. The multimodal distributions of *Ecrins* and *Canjuers* subsets (Figure 7) could reflect such
bias, potentially exacerbated by the lower number of their removals (n = 12 and n = 11) regarding
other subsets.

622 **4.3 Causality inference**

623 Potential ripple effects of removals raise the more general question of the contribution of removals 624 in the trends we observed. As the variability of subsets showed, the attack process could seem 625 inconstant, with intensity irregularly turning on and off at the same spatial distances even before 626 removals. If contribution of removals in changing the dynamics and intensities of attacks at low spatial and temporal scales could be easily considered, our confidence in such causal relationship 627 628 was decreasing as analysis scale increased. For example, in the subset of Non-interacting removals 629 (Figure 6.A), is the disappearance, after removals, of the patch of high attack intensity between 6 630 and 8 km from removals present at -60 days before removals the result of removals? Apart from 631 livestock presence, many confounding factors can also affect attack intensity distribution, starting 632 with wolves themselves, e.q. their varying density. If dividing removals into seasonal subsets may 633 smoothen differences between removals due to the wolf annual biological cycle, they could not 634 entirely erase random biological events that rhythmed wolf lives, as dispersal, settlement, or natural 635 deaths. Restricting the analysis at scale 1 could greatly reduce such bias. As an example, our 636 correction for livestock presence greatly impacted results at scale 3 but not at scale 1. However, 637 because wolves move over large territories, we acknowledge that focusing on changes happening 638 only in close proximity to removals could not be enough to evaluate removal effects (Treves et al., 639 2016).

640 **4.4 Exploring variability in the spatio-temporal distribution of attacks**

641 Keeping these limits in mind, we are confident that our results showed that there could not be a 642 single type of changes in attack intensities after lethal control, but several, and that they form a 643 gradient noticeable in the heatmaps. First, after lethal control, attack intensity could decrease on the 644 long term (observed results of the subsets of Préalpes de Digne, or of Trois-Evêchés) or on the short 645 term only (Denning). Second, attack intensity could decrease but not immediately after lethal 646 control (Hunting, Female, Fall, Reproduction, Mercantour). Third, the duration of high intensity 647 patches (Simple defense) or their spatial expansion (Non-interacting, PNRPA, Male, Fall) or both at 648 the same time (Interacting, Ecrins), could decrease after lethal control. Fourth, no major changes in 649 attack intensity could be observed after lethal control even on the long term (*Reinforced defense*, 650 Vanoise, Canjuers). Fifth, the spatial expansion of high intensity patches could increase (Rendez-651 vous). Sixth, the attack intensity could be sliced into several patches of high intensity increasing with

652 time (Vercors, Winter).

The origins of these variations between subsets were hard to explain. They could originate from 653 654 pastoral practices, including non-lethal measures, combined to other environmental differences 655 such as altitude that can modify wild ungulate distributions or wolf distribution (Miller et al., 2016). 656 The choices we made to define subsets were an attempt to group removals applied in similar 657 conditions or giving similar results (sex of killed wolf), expecting common patterns of attack intensity 658 distribution would emerge. Ideally, we could have grouped removals according to location and 659 season (and even years) at the same time, but we were limited by the number of removals per subset. On the contrary, pooling a large number of removals, as in the Non-interacting subset, may 660 661 be counterproductive because a large subset would include situations that are too various to lead to

similar effects of lethal control, therefore preventing patterns in attack intensity distribution fromemerging.

664 Some hypotheses concerning the variability of results could be made. They rest upon consideration 665 of local and temporal context. For example, at scales 1 and 2, the removals from the Simple defense 666 subset appeared more efficient than those of the Reinforced defense subset to reduce attack 667 intensities (Table 1, Figures 6.I, 6.J). This result could originate from the fact that reinforced defense 668 removals were allowed when the protected flocks, or other flocks from the same municipality, have 669 recurrent attacks despite simple defense removals (Ministère de la Transition Ecologique et Solidaire, 2018). In other words, reinforced defense removals were supposed to be applied in 670 671 situations that were not mitigated through simple defense removals, maybe because of a higher 672 wolf density. However, the removals of the Hunting subset, supposed to be applied in situations that 673 were not mitigated through both defense removals, were associated with a decrease of attack 674 intensities that did not appear immediately after removals but that was long-lasting. This result 675 could originate from the fact that hunting removals occurred essentially in October, contrary to the defense removals mainly occurring in August. It could have been easier, in the winter context, to 676 677 maintain attack intensity at the level they were expected (e.g. flocks in fold during the night). In 678 addition, the analyzed reinforced defense removals were mostly applied in the last years of the 679 study period (58% in 2019-2020), whereas the hunting removals were mostly applied in 2015-2016 680 (51%). Estimated wolf population size increased greatly between these two years, being 341 [95% IC

681 266-463] in 2015 and 624 [95% IC 414-834] in 2020 (Réseau Loup-Lynx, 2021).

- 682 Thus, reducing attack intensity in summer whereas wolf population was densifying appeared a
- 683 difficult task for reinforced defense removals. Reducing locally and temporarily wolf density could

have improved lethal control efficiency, as found by Bradley et al. (2015). In our results, killing more

than one wolf in less than 5 km and less than 25 days, as interacting removals did, indeed appeared

686 more efficient to reduce attack intensities than killing only one wolf, even if patches of high intensity

687 still persisted after interacting removals (Figure 6.A-6.B, Table 1).

688 4.5 Consistency with previous studies

689 Our most common result of a moderate and temporary decrease of attack intensity following lethal

690 control, tempered by a large variability depending on contexts and on analysis scales, was consistent

691 with most of the previous comparative studies (*e.g.* Bradley et al., 2015; Harper et al., 2008;

Santiago-Avila et al., 2018) and the apparent discrepancy on the subject (Lennox et al., 2018; Treveset al., 2016).

Some disparity with previous studies also appeared. Contrary to Bjorge and Gunson (1985) or
Bradley et al. (2015) who reported total or high efficiency of lethal control on reducing attacks in

time, the time duration effectiveness in our results was generally much lower. However, very

697 different management contexts applied in North America where these studies were conducted

698 compared to the French context. Our most intense removals (on average 2.3 wolves for interacting

removals) corresponded to the least intense in these studies (*e.g.* 2.2 wolves for partial removal in

700 Bradley et al. (2015)). In addition, the Western Alps are close to saturation in wolf packs (Marucco et

al., 2023). Therefore, we cannot expect the same recurrence of attacks nor the same effectiveness in

702 lethal control in both situations.

703 Compared to the previous studies, the strength of our analysis was to demonstrate the variability of

704 lethal control effects on both continuous spatial and temporal scales by computing attack intensities

705 through kernel density estimates. Last but not least, the control for livestock presence, even if

706 perfectible with more updated data, reduced the risk of confounding factors and improved inference 707 of this non-experimental study. For some subsets as *Winter*, control for livestock presence drastically

708 changed the results.

709 4.6 Management implications

- 710 For conditions similar to the application of lethal control in France, managers should keep in mind
- that applying lethal removals may reduce attack intensities at small scale (i.e. below 2 km and 15
- 712 days), but with several drawbacks. First, attacks are likely to persist after removals, including at high
- intensity. Second, absence of changes or even increase of attack intensities after removals are still
- possible. Third, trends in attack intensity are more variable at larger scales. In other words, lethal
- 715 control can contribute to management of wolf attacks but cannot be expected to cancel by itself
- 716 predation on livestock on large timescales or on large areas with removal intensities as those
- analyzed, even if its implementation goes along with non-lethal measures as in the Western Alps.
- 718 In addition, the high variability and context-specificity of our results support a management which is
- 719 itself contextualized, but the lack of understanding of this variability prevents us to give more
- 720 specific management implications per context. Our results cannot be considered as predictions of
- 721 lethal control effects on attack distribution, even per subset, as too many factors influence the
- 722 attack process beyond the characteristic of the subset and may evolve.
- 723 Moreover, managers cannot easily handle the unit of measure of intensity we used, *i.e.* the number
- of attacks over a small spatio-temporal area (50 meters 1 day) corrected for spatial bias and for
- 725 livestock presence. If managers can do so with intensity trends from Table 1, we strongly caution,
- however, to read Table 1 without the corresponding heatmaps. Indeed, the heatmaps provided the
- 727 detailed range of changes of the attack intensity distribution after removals, and therefore the
- potential responses of wolves to removals in its full complexity. On the contrary, Table 1 could only
- 729 indicate three tendencies (decrease, null, increase) that tended to oversimplify lethal control effects
- py concealing other components of the evaluation (e.g. frequency of intense attack events, duration,
- 731 potential attack shifts...).

732 4.7 Perspectives

- 733 Evaluation of the effects of lethal control on attacks made by large carnivores on livestock remains a
- technical challenge. Our results showed that consideration of analysis scale is crucial and that effects
- should be analyzed separately for each local context.
- 736 To fully understand the effects of lethal control, we need more precise and long-term data related to
- 737 carnivores for example about their age, their pack affiliations or their reproductive status, through
- 738 genetics or telemetry. We also need up-to-date information about livestock presence and more
- 739 precise data about pastoral practices (*e.g.* livestock non-lethal protection). Collection of these
- 740 datasets and their analysis are particularly feasible at small scale, for example at the scale of one
- 741 wolf territory and their neighbors. In addition, small scale analysis is consistent with the need to
- contextualize evaluation of lethal removal effects. Effects of lethal control on wolves and their
- attacks may then be evaluated through promising methods such as animal social network analysis,
- which was applied to study the resilience of other social animals to human-induced removals (*e.g.*
- 745 Downing et al., 2023; Goldenberg et al., 2016).
- The side effect of focusing on a specific context is to restrict the number of cases and to fail to
- 747 extract a general pattern if any. A solution would be to upgrade the analysis of observed data by
- 748 simulations. For example, Wiśniewska et al. (2022) used individual-based modelling to simulate

- poaching effects on elephant social structures. This type of modelling is useful to reproduce
- biological processes even for socially complex species as wolves. Haight et al. (2002) already used
- this method to simulate management of attacks of wolves by lethal control, but it lacked several
- biological processes that have been implemented in new population models since that could be used

753 for this purpose (*e.g.* Bauduin et al., 2020).

- 754 Thus, combination of analysis based on observed data as social network analysis and simulations
- through individual based-models may be the next step to evaluate lethal control effects on social
- carnivores while understanding the biological responses of the targeted species to lethal control.

- . . .

782 Appendices



784Figure S1. Results of kernel estimation of attack intensities surrounding lethal removals, detailed for each analyzed subset785of removals. Each panel includes (i) the map of the removal locations with the annotated number of involved removals, (ii)786the monthly distribution of the removal dates, (iii) the observed attack intensities λ_{obs} and (iv) the corrected attack787intensities λ_{corr} . Intensities are given along the continuous temporal axis (x-axis) and spatial axis (y-axis). The coordinate788(0,0) indicates the date and the location at which removals occurred. Grey rectangles of subpanels iii and iv delimit the789three analysis scales 1 (0-2 km, ± 15 days), 2 (0-6 km, ± 45 days) and 3 (0-10 km, ± 90 days). A value of λ_{corr} lower, equal or790higher than zero indicated that the observed attack intensity was lower, equal or higher, respectively, than the attack



D. Mercantour









Novembe

August

July

Number of lethal removals

October

September

February

May

June

March

April









G. Reproduction







ii.



H. Lethal removal of male





- 797 Figure S1. (Part 4/4)
- 798

- 799
- 800

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809 Data, scripts and code

- 810 Data, scripts and code are available online: <u>https://doi.org/10.5281/zenodo.12772868</u> and
- 811 <u>https://gitlab.com/oksanagrente/Kernel_wolf_culling_attacks_p</u>

812 Conflict of interest disclosure

- 813 The authors declare that they comply with the PCI rule of having no financial conflicts of interest in
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- 986



Removal type Simple defense Reinforced defense Hunting Interacting

Pastoral surfaces Attacked Unattacked



| Unused pastoral surface | | First day of m ₃ | | Last day of m ₃ |
|---------------------------------------|--------|---------------------------------|-------------------------------|------------------------------------|
| 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 | Ring r | ра _{г,т3,75} | ра _{г, тз,} | ра _{г,т3,90} |
| | 1 | ра _{1,т3,75} | ра _{1,т3,} | pa _{1,m3,90} |
| | 2 | ра _{2,т3,75} | ра _{2,т3,} | ра _{2,т3,90} |
| | | ра _{,m_{3,75}} | ра _{,m_{3,}} | ра _{,т_{3,90}} |
| | 199 | ра _{199,т3,75} | ра _{199, тз,} | ра _{199,т_{3,90}} |
| | 200 | ра _{200,m3,75} | ра _{200, тз,} | ра _{200,т3,90} |





Altitude (m)

| 0 |
|------|
| 1000 |
| 2000 |
| 3000 |
| 4000 |

Season

| | Winter |
|------------------|--------------------|
| | Downoduction |
| | Reproduction |
| \bigtriangleup | Den |
| | Rendez-vous |
| | Fall |
| | |

Zone

| Attributed |
|----------------|
| Not attributed |



A. Non-interacting lethal removals



November



B. Interacting lethal removals





Number of lethal removals





Temporal distance to removals (day)

C. PNRPA





D. Préalpes de Digne



E. Vercors

F. Den

G. Rendez-vous

H. Winter

0.25

0.00

-0.25

-0.50

-0.75

0.6

0.4

0.2

I. Simple defense

J. Reinforced defense

K. Hunting

