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V. Squicciarini, J. Mazoyer, A. -M. Lagrange, A. Chomez, P. Delorme, et al.. The COBREX archival survey: improved constraints on the occurrence rate of wide-orbit substellar companions. I. A uniform re-analysis of 400 stars from the GPIES survey. Astronomy & Astrophysics - A&A, In press. hal-04778163

HAL Id: hal-04778163 https://hal.science/hal-04778163v1

Submitted on 12 Nov 2024 $\,$

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The COBREX archival survey: improved constraints on the occurrence rate of wide-orbit substellar companions

I. A uniform re-analysis of 400 stars from the GPIES survey

V. Squicciarini¹, J. Mazoyer¹, A.-M. Lagrange^{1,2}, A. Chomez^{1,2}, P. Delorme², O. Flasseur³, and F. Kiefer¹

¹ LESIA, Observatoire de Paris, Université PSL, CNRS, Sorbonne Université, Université Paris Cité, 5 place Jules Janssen, 92195 Meudon, France

e-mail: vito.squicciarini@obspm.fr

- ² Univ. Grenoble Alpes, CNRS-INSU, Institut de Planétologie et d'Astrophysique de Grenoble (IPAG) UMR 5274, Grenoble, F-38041, France
- ³ Centre de Recherche Astrophysique de Lyon (CRAL) UMR 5574, CNRS, Univ. de Lyon, Univ. Claude Bernard Lyon 1, ENS de Lyon, F-69230 Saint-Genis-Laval, France

November 8, 2024

ABSTRACT

Context. Direct imaging (DI) campaigns are uniquely suited to probing the outer regions around young stars and looking for giant exoplanet and brown dwarf companions, hence providing key complementary information to radial velocity (RV) and transit searches for the purpose of demographic studies. However, the critical 5-20 au region, where most giant planets are thought to form, remains poorly explored, lying in-between RV and DI capabilities.

Aims. Significant gains in detection performances can be attained at no instrumental cost by means of advanced post-processing techniques. In the context of the COBREX project, we have assembled the largest collection of archival DI observations to date in order to undertake a large and uniform re-analysis. In particular, this paper details the re-analysis of 400 stars from the GPIES survey operated at GPI@Gemini South.

Methods. Following the pre-reduction of raw frames, GPI data cubes were processed by means of the PACO algorithm. Candidates were identified and vetted based on multi-epoch proper motion analysis – whenever possible – and by means of a suitable color-magnitude diagram. The conversion of detection limits into detectability maps allowed for an estimate of unbiased occurrence frequencies of giant planets and brown dwarfs.

Results. Deeper detection limits were derived compared to the literature, with up to a twofold gain in minimum detectable mass compared to previous GPI-based publications. Although no new substellar companion was confirmed, we identified two interesting planet candidates awaiting follow-up observations. We derive an occurrence rate of $1.7^{+0.9}_{-0.7}\%$ for 5 $M_{Jup} < m < 13 M_{Jup}$ planets in 10 au < a < 100 au, that raises to $2.2^{+0.8}_{-0.8}\%$ when including substellar objects up to 80 M_{Jup} . Our results are in line with the literature, but come with lower uncertainties thanks to the enhanced detection sensitivity. We confirm the finding, hinted at by previous studies, of a larger occurrence of giant planets around BA hosts compared to FGK stars; moreover, we tentatively observe a smaller occurrence of brown dwarf companions around BA stars, although larger samples are needed to shed light on this point.

Conclusions. While waiting for the wealth of data expected from future instrument and facilities, valuable information can still be extracted from existing data. In this regard, a complete re-analysis of SPHERE and GPI data is expected to provide the most precise demographic constraints ever provided by imaging.

Key words. Planets and satellites: detection – Planets and satellites: gaseous planets – brown dwarfs – Planetary systems – Techniques: high angular resolution

1 1. Introduction

Bolstered by more than 5000 confirmed detections to date, the 2 exoplanet field has become mature enough to accompany the 3 still thriving detection-oriented endeavor with follow-up stud-4 ies aimed at shedding light on key questions related to the ori-5 gin, the prevalence, and the architecture of planetary systems. 6 By unveiling statistical trends in the measured physical, orbital 7 and star-related properties of the exoplanet population, exoplanet 8 demographics seeks to connect theory and observation, in order 9 to fully understanding the physical processes underlying planet 10 formation (Biazzo et al. 2022). 11

The census of known exoplanets currently spans about four 12 magnitudes in mass and in semi-major axis¹. No single detection 13 method is adequate to probe such a large extent of the parame-14 ter space: it is through the combination of the different meth-15 ods, each optimized for detection inside a specific niche, that 16 the large-scale picture can be unveiled and reconstructed (see, 17 e.g., Gratton et al. 2023, 2024). However, obtaining a complete 18 and unbiased blend from heterogeneous ingredients is hindered 19 by factors such as inconsistent detection criteria, completeness 20 and false positive assessment, uncertainty quantification, neglect 21 of underlying selection or observational biases (Gaudi et al. 22

¹ Empirically estimated based on the Extrasolar Planet Encyclopaedia: http://www.exoplanet.eu/.

2021). Whenever two different methods can be simultaneously 23 employed, their complementarity allows better characterizing in-24 dividual objects (see, e.g., Gandolfi et al. 2017; Bonnefoy et al. 25 2018; Bourrier et al. 2018; Lacedelli et al. 2021; Kuzuhara et al. 26 2022; Philipot et al. 2023, Lagrange et al. under review) and 27 strengthening the statistical trends emerging in each of the meth-28 ods (Rogers 2015; Santerne et al. 2016). In the cases where dif-29 ferent techniques probe instead different separations within the 30 same system, the joint analysis opens up exquisite dynamical 31 and formation studies (see, e.g., Covino et al. 2013; Bryan et al. 32 2016; Zhu & Wu 2018). 33

Radial velocity (RV) surveys have provided invaluable con-34 straints on the physical and orbital properties of giant planets 35 up to ~ 5 au (Wolthoff et al. 2022; Rosenthal et al. 2024). Yet, 36 the reliability of RV trends for larger separations has been ques-37 tioned (Lagrange et al. 2023), and the predicted yields for di-38 39 rect imaging (DI) surveys based on extrapolations of RV results 40 have been shown to be too optimistic (see, e.g., Cumming et al. 2008; Dulz et al. 2020). On the other hand, direct imaging (DI) 41 is mostly sensitive to young giant planets in wide ($a \ge 20$ au) 42 orbits, providing access to the scarcely studied outskirts of plan-43 etary systems. Starting from 2004 (Chauvin et al. 2004), direct 44 imaging has discovered ~ 30 planets ($M < 13M_{Jup}$) (Zurlo 45 2024), including iconic systems like the disk-enshrouded PDS 46 70 (Keppler et al. 2018), the ~ 20-Myr-old β Pictoris (Lagrange 47 et al. 2009), 51 Eridani (Macintosh et al. 2015), AF Leporis 48 (Mesa et al. 2023; De Rosa et al. 2023; Franson et al. 2023), 49 and the four-planet HR 8799 (Marois et al. 2008). These de-50 tections are the main outcome of large blind surveys targeting 51 52 tens (e.g. MASSIVE, Lannier et al. 2016; SEEDS, Uyama et al. 53 2017; LEECH, Stone et al. 2018) or hundreds of stars (e.g. NICI-PCF, Liu et al. 2010; IDPS, Galicher et al. 2016; ISPY-NACO, 54 Launhardt et al. 2020). The forefront of DI surveys, enabled by 55 the exquisite performances of imagers and integral field spec-56 trographs, coupled with extreme AO systems mounted on 8-m-57 class telescopes, is currently represented by the 400-star SHINE 58 (Chauvin et al. 2017) and the 600-star GPIES (Nielsen et al. 59 2019) surveys. 60

By constraining the overall frequency and the properties of 61 wide-separation giant planets, DI studies are expected to enable 62 a thorough comparison with concurrent formation models (see, 63 e.g., Bowler 2016; Vigan et al. 2021); orbital properties, for in-64 stance, shed light upon their formation and dynamical evolution 65 (Bowler et al. 2020); the dependence of frequency on stellar 66 mass provides clues about the initial state of the disk and the 67 formation mechanisms at play (Nielsen et al. 2019; Janson et al. 68 2021). However, despite years of extensive searches, it is still 69 not clear whether the main formation channel for the observed 70 wide-orbit population be core accretion (CA; Pollack et al. 1996; 71 Mordasini et al. 2009), the bottom-up process responsible for the 72 formation of planets in the Solar System, or rather a top-down 73 star-like scenario like gravitational instability (GI; Boss 1997; 74 Vorobyov 2013). While an interplay between the two scenarios is 75 deemed to be favored by empirical parametric models (Reggiani 76 et al. 2016; Vigan et al. 2021) and direct comparison with syn-77 thetic planet populations (Vigan et al. 2021) alike, understanding 78 in an unambiguous way how each known companion was formed 79 is still beyond reach. The large uncertainties still existing in the 80 interpretation of the observed picture can be attributed, at least 81 partially, to the fact that the critical 5-20 au region, where most 82 giant planets are thought to form, remains poorly explored being 83 exactly in-between current RV and DI capabilities. 84

Under given observing conditions, the final performances attainable by a high-contrast imaging observation are dictated both by instrumental (e.g., the telescope, the science instru-87 ment, the performance of adaptive optics and coronagraphs) and 88 post-processing components (the algorithms applied to science 89 images to decrease the level of systematic and random noise) 90 (Galicher & Mazoyer 2024). Depending on observing condi-91 tions, stellar brightness and angular separation, state-of-the-art 92 instruments such as the Spectro-Polarimetric High-Contrast Ex-93 oplanet Research (SPHERE; Beuzit et al. 2019) and the Gem-94 ini Planet Imager (GPI; Macintosh et al. 2014) typically achieve 95 raw planet-to-star contrasts as low as $10^{-3} - 10^{-5}$ (Poyneer et al. 96 2016; Courtney-Barrer et al. 2023). On the instrumental side, 30-97 m-class telescopes and space-borne coronagraphic instruments 98 are expected to bring about a major leap forward for the field in 99 the next decade (see, e.g., Kasdin et al. 2020; Kasper et al. 2021). 100 whereas upgrades of existing instruments such as SPHERE+ 101 (Boccaletti et al. 2022) and GPI 2.0 (Chilcote et al. 2018) are 102 going to represent the forefront in the medium term; on the re-103 duction side, advanced post-processing algorithms have been al-104 ready shown to increase the contrast by as much as two orders 105 of magnitudes compared to pre-reduced data. Therefore, the de-106 velopments of more powerful reduction techniques can greatly 107 increase detection capabilities working on observations that al-108 ready exist (see, e.g., Currie et al. 2023). 109

In the framework of the COupling data and techniques 110 for BReakthroughs in EXoplanetary systems exploration (CO-111 BREX) project, we collected more than a thousand archival 112 SPHERE and GPI observations, assembling the largest exoplan-113 etary direct imaging survey to date, with the aim of re-reducing 114 them in a uniform and self-consistent way. The results of the 115 full re-reduction of the SHINE survey are illustrated in Chomez 116 et al. (2024). In this work, we present the re-reduction of 400 117 stars coming from GPIES. Despite being the largest DI obser-118 vational campaign to date, just two new substellar objects were 119 discovered during the survey: one planet (51 Eri b, Macintosh 120 et al. 2015) and one brown dwarf (HR 2562 B, Konopacky et al. 121 2016). A statistical analysis of the first 300 stars was performed 122 by Nielsen et al. (2019, hereafter N19). By combining the two 123 surveys, it will be possible to obtain the tightest constraints to 124 date on the occurrence of wide-orbit giant planets, hence provid-125 ing an ideal test-bed to scrutinize planet formation models. 126

This paper is organized as follows: after laying out the se-127 lection criteria for the sample and the corresponding observa-128 tions (Section 2.2), and uniformly deriving the stellar parame-129 ters of interest (Section 2.3), we describe in detail the process of 130 data reduction (Section 2.4). Section 3 presents the results of the 131 analysis, namely companion candidates and completeness maps. 132 The derived occurrence rates are presented and discussed in Sec-133 tion 3.4. A thorough comparison with the literature is the subject 134 of Section 4. Finally, in Section 5, we summarize the results of 135 this work. 136

2. Data

2.1. Raw data collection

The observations considered in this work were collected between 139 2013 and 2020 by means of GPI at the Gemini South telescope. 140 GPI is an integral-field spectrograph (IFS) with low spectral res-141 olution (~50; Maire et al. 2014), operating in the wavelength 142 range [0.97-2.40] μ m. As the vast majority of GPI observations 143 were gathered in the H band ([1.5–1.8] μ m) (Ruffio et al. 2017) 144 - other bands being mostly used for characterization purposes -145 we decided to restrict our query to H-band observations. 146

137

We downloaded therefore all H-band raw frames from GPI 147 that are publicly available on the Gemini archive² (~30000 148 $(frames)^3$. We neglected observational sequences with 5 or less 149 frames (typically corresponding to \leq 5-min exposure times), and 150 observations of stars only taken for calibration purposes (easily 151 identifiable through their program ID). The reason behind this 152 choice is twofold: on the one hand, the known multiplicity of 153 these stars is expected to detrimentally affect the performances 154 attainable by post-processing; on the other hand, selection crite-155 ria for these stars are different from those of science targets, thus 156 inducing a bias when interested in statistical considerations. 157

The preliminary sample obtained in this way (hereafter GPI 158 database) is composed of 852 sequences for 655 stars. Most of 159 the observations (715/852) within the database were collected in 160 the course of GPIES, and additional 10 sequences are describ-161 able as follow-up observations of interesting stars from the cam-162 paign. Intertwined to GPIES observations, the remaining 127 se-163 quences were gathered over the lifetime of the instrument as part 164 of other scientific programs. The following Section 2.2 will elu-165 cidate how the final stellar sample was assembled, and the cri-166 teria that a sequence had to meet in order to be included in the 167 corresponding sample of observations. 168

169 2.2. Sample definition

Like any other direct imaging search to date, GPIES was con-170 171 structed by looking for young stars in the solar neighborhood; the reason lies in the fact that recently formed exoplanets and 172 brown dwarfs are brighter and hotter than mature objects of 173 the same mass due to residual formation heat, yielding a sig-174 nificantly more favorable planet-to-star contrast in near infrared 175 bands. In particular, the stellar sample was assembled by merg-176 177 ing lists of members of young moving groups from the literature (de Zeeuw et al. 1999; Zuckerman et al. 2001, 2011) with close 178 (<100 pc) stars selected for large X-ray emission. Echelle spec-179 180 tra were obtained for ~ 2000 stars to further identify additional 181 young stars based on lithium abundance and chromospheric ac-182 tivity (see Section 2.3 for details). After removing apparent binaries with angular separation $\in [0.02^{\circ}, 3^{\circ}]$ and $\Delta mag < 5$ (both 183 before and during the campaign), and accounting for some new 184 association members, a sample of 602 stars was finally obtained 185 (Nielsen et al. 2019). 186

Due to the decommissioning of GPI – currently undertaking 187 major upgrades to become GPI 2.0 (Chilcote et al. 2020) - in 188 early 2020, the GPIES survey was never completed (~ 10% of 189 the stars lack observations). It is thus vital to ascertain whether 190 the sample of observed stars be an unbiased extraction of the full 191 192 sample. An automated target-picker was employed to suggest the best targets for every telescope night, as a function of both 193 observing conditions and stellar age/distance (McBride et al. 194 2011); however, we verified through a Kolmogorov-Smirnov test 195 $(\alpha = 0.05)$ that the age and distance distributions of the first 300 196 stars (those from N19) are compatible with those of the full sam-197 ple of observed stars ($p_{age} \approx 1$, $p_{dist} = 0.051$). Additionally, stars 198 with known companions were not prioritized in their first-epoch 199 observation (Nielsen et al. 2019). We can thus confidently main-200 tain that the available GPIES observations are not affected by 201

selection biases, making the sample suitable for statistical studies. 202

The definition of the final sample was based on a combina-204 tion of observational constraints and physical constraints on stel-205 lar properties. As regards the former aspect, a minimum amount 206 of parallactic angle rotation $\Delta PA \sim 10^{\circ}$ is required to enable 207 efficiently using angular differential imaging (ADI) during post-208 processing (Marois et al. 2006): we conservatively adopt a mini-209 mum rotation of 12° in order to exploit angular differential imag-210 ing (ADI) during post-processing⁴. A single observation with 211 extremely bad seeing was removed. With respect to the latter, 212 we only retained stars for which youth can be established with 213 reasonable confidence (see Section 2.3). 214

Given our ignorance about the selection criteria adopted 215 for stars from non-GPIES programs, we decided not to consider them for the purpose of this paper. However, we retained 217 non-GPIES observations of GPIES stars as valuable follow-up 218 epochs for promising point-source candidates. The final sample 219 employed throughout this work consists of 400 stars (515 sequences). 221

2.3. Stellar parameters

The knowledge of stellar ages is pivotal to a meaningful interpretation of direct imaging campaigns, as a large degeneracy exists between age and mass – let alone additional parameters like metallicity or a planet's formation history – for substellar objects for which only photometric data are available (see, e.g., Spiegel & Burrows 2012).

Our primary age diagnostics is provided by kinematic mem-229 bership to young associations and moving groups (hereafter 230 YMGs). Starting from Gaia DR3 (hereafter Gaia; Gaia Collab-231 oration et al. 2023) data, we used BANYAN Σ (Gagné et al. 232 2018b) to classify a star as a member of a YMG if the associ-233 ated membership probability p > 90%. A second indicator was 234 represented by the ages obtained by N19: we stress that the un-235 derlying data are not public and the derived ages, not equipped 236 with error bars, are only available for the 300 stars presented in 237 that study. Finally, the ages for 21 additional stars – that are not 238 members of YMG nor targets of N19 – could be recovered after 239 cross-matching our sample with SHINE (Desidera et al. 2024), 240 whose thorough analysis builds upon a manifold variety of in-241 dicators (isochrones, YMG membership, activity, lithium abun-242 dance). For ages based on N19, that come with no associated 243 uncertainty, we adopt a constant fractional uncertainty of 25%, 244 empirically tuned to match the typical fractional uncertainty for 245 SHINE stars. 246

Individual stellar parameters for each star were obtained by 247 means of MADYS⁵ (v1.2, Squicciarini & Bonavita 2022), a tool for 248 (sub)stellar parameter determination based on the comparison 249 between photometric measurements and isochrone grids derived 250 from theoretical (sub)stellar models. Assuming the ages de-251 scribed above, photometry from Gaia DR3 and 2MASS (Skrut-252 skie et al. 2006) – corrected by extinction by integrating the 3D 253 map by Leike et al. (2020) along the line of sight – was com-254 pared to the last version of non-rotating, solar-metallicity PAR-255 SEC isochrones (Nguyen et al. 2022). 256

The distance, age and spectral type of the 400 stars are shown 257 in Figure 1. The full collection of the derived properties is sum-258

Article number, page 3 of 21

² https://www.cadc-ccda.hia-iha.nrccnrc.gc.ca/en/ gemini/ and https://archive.gemini.edu/

³ We could not retrieve science commissioning data from 2014 from the old Gemini website (Macintosh et al. 2014). However, the targeted stars were only known hosts of exoplanets or disks (HR 8799, HR 4796, HD95806) that were later re-observed during GPIES.

⁴ The distribution of ΔPA across the database shows a bump for $\Delta PA > 12$. Whereas just 12 sequences have $10^{\circ} < \Delta PA \le 12^{\circ}$, 49 observations are found inside the $12^{\circ} < \Delta PA \le 14^{\circ}$ bin.

⁵ https://github.com/vsquicciarini/madys



Fig. 1: Age of the final stellar sample as a function of distance. The color scale labels different spectral types. Kernel density estimates for the distribution of the two properties are provided on top and to the right of the main plot.

marized in Table A, while the ages adopted for YMGs are provided in Table A.2.

261 2.4. Data reduction

262 2.4.1. Preliminary steps

The pre-reduction of raw GPI data is performed in two steps: the goal of the first step is to build a 3D data cube (x, y, λ) starting from the 2D image acquired in the detector plane; the collection of the derived data cubes is then stacked into a final 4D data cube (x, y, λ, t) .

As routinely done for GPI data, we employed the version 268 1.6.0 of the GPI Data Reduction Pipeline (hereafter DRP; Perrin 269 et al. 2014, 2016; Wang et al. 2018) to handle the first part of 270 the pre-reduction. After subtracting a dark frame, bad pixels in 271 the image are substituted by interpolated values. The mapping 272 of the \sim 37000 small spectra created on the detector by GPI's 273 lenslet array to the corresponding spaxels of the 3D data cube is 274 determined by means of master wavelength calibrations based on 275 Xenon or Argon lamps, conveniently corrected for mechanical 276 offsets induced by flexure (Wolff et al. 2014). The signal of each 277 spectrum can be thus extracted and stored into the corresponding 278 spaxel (Draper et al. 2014). Subsequent steps correct for small 279 distortion effects of the field of view and for halos induced by 280 residual atmospheric turbulence. 281

Thanks to a square grid embedded within the pattern of 282 the apodizer, a diffraction pattern of satellite spots (hereafter 283 satspots) – attenuated images of the star – is created in the image. 284 The four first-order satspots, symmetrically situated at ~ $20\lambda/D$ 285 from the star, serve three different purposes: 1) to recenter the 286 frames, by locating the position of the occulted star; 2) to cali-287 brate the flux level of each pixel in the science image; 3) to build 288 a model of the off-axis point spread function (PSF)⁶ (Wang et al. 289 2014). The final operations of the DRP deal with the astrometric 290 and photometric characterization of the satellite spots, fitted by 291 a Gaussian PSF template. 292

The second step of the pre-reduction deals with stacking up 293 and recentering frames to build the final 4D data cube. In ad-294 dition to the data cube, three files are created: 1) a 4D PSF: 2) 295 a wavelength vector; 3) a parallactic angle vector, that indicates 296 the rotation of the field of view during the sequence. Indeed, GPI 297 observations are operated in a pupil-stabilized mode, i.e. with no 298 derotator, so as to allow using ADI-based post-processing algo-299 rithms (see, e.g., Ruffio et al. 2017). 300

This stage is achieved using PYKLIP (Wang et al. 2015a). 301 Compared to the 2.6 version, we introduced slight modifica-302 tions⁷ to create output files formatted in a SPHERE-like way as 303 regards the data format, the PSF, and data cube flux, and the FOV 304 orientation (East to the left). In this way, we ensured the har-305 monization of the future SPHERE+GPI sample while smooth-306 ing the I/O integration with the post-processing algorithm (Sec-307 tion 2.4.2). In addition to computing the image center through 308 satspot pattern, PYKLIP estimates satspot flux in a more precise 309 way than the DRP, a crucial step for photometric characteriza-310 tion purposes, and empirically recomputes the wavelength vec-311 tor based on the satspot-to-center separation (which scales with 312 λ). We visually checked the goodness of the result for all our im-313 ages; whenever a specific satspot was - due to intrinsic dimness 314 or systematic problems – not properly fitted, inducing centering 315 offsets in one or more frames, we used a specific option of PYKLIP 316 to ignore it during recentering. 317

As already mentioned, satellite spots are faint images of the target star; the flux ratio between a satspot and the star, or grid ratio, was determined by Wang et al. (2014) to be $\Delta m = 9.4 \pm 0.1$ 320 mag through on-sky observations. The ~ 10% uncertainty on the grid ratio turns out to be one of the main factors in the total error budget of GPI spectrophotometry. 323

We performed several tests to quantify the reliability of the 324 wavelength solution, the image centering and the photometric 325 calibration of the PSF. The accuracy of the DRP wavelength so-326 lution was estimated by Wolff et al. (2014) to be 0.032% in H-327 band, well below the 1% accuracy needed to achieve a spectral 328 characterization uncertainty < 5%. As regards the wavelength 329 precision, we collected for every sequence the satspot positions 330 estimated by the DRP, computed separations from PYKLIP's frame 331 centers, $\xi(\lambda, t, s)$, then averaged over the temporal axis t and the 332 satspot axis s to obtain $\hat{\xi}(\lambda)$. We computed the ratio $\eta = \xi(\lambda)/\lambda$ 333 for every sequence, a value that ought to be constant. The 50th, 334 16th and 84th percentiles of the distribution across the sequences 335 yields $<\eta>=37.921^{+0.056}_{-0.046}$ px/ μ m, corresponding to a precision 336 of 0.15%. 337

As regards centering precision, propagation of random un-338 certainties in PYKLIP yields a centering precision along one axis 339 $\sigma_{c,x} = 0.04$ px (that is, ~ 0.06 px in 2D), comparable to the one 340 reported in Wang et al. (2014). However, we identified a system-341 atic deviation of the satspot pattern shape from a square (which is 342 an underlying assumption of PYKLIP's centering algorithm): the 343 difference between the centers computed from doublets of op-344 posite satspots $(x_{c,13}, y_{c,13})$ and $(x_{c,24}, y_{c,24})$, stable over time, is 345 $\Delta S = \sqrt{(x_{c,13} - x_{c,24})^2 + (y_{c,13} - y_{c,24})^2} = 0.39 \pm 0.04$ px. We 346 consider the true center to be distributed according to a uniform 347 distribution between $(x_{c,13}, y_{c,13})$ and $(x_{c,24}, y_{c,24})$. The final cen-348 tering error σ_c is therefore: 349

$$\sigma_c = \sqrt{\left(\frac{\Delta S}{\sqrt{12}}\right)^2 + 4 \cdot (\sigma_{c,x})^2} \approx 0.15 \text{ px} \sim 2 \text{ mas.}$$
(1)

⁶ Unlike SPHERE, no PSF exposure is taken before and/or after the scientific observation in GPI.

⁷ url of GitHub repository, to be added upon acceptance

This value was consistently employed when propagating astrometric uncertainties of detected sources.

Finally, we adopted a platescale of 14.161 ± 0.021 mas px⁻¹ (De Rosa et al. 2020a), assuming it to be stable over time (Tran et al. 2016). With respect to the north offset angle, we used a time-varying value following the prescriptions indicated in Table 4 from De Rosa et al. (2020a).

357 2.4.2. Post processing - PACO

Pre-reduced datasets were processed in the COBREX Data 358 Center, an improved version of the High-Contrast Data Cen-359 ter (HC-DC⁸, formerly SPHERE Data Center, Delorme et al. 360 2017). Prompted by the promising preliminary results presented 361 in Chomez et al. (2023), we decided to process our archive 362 by means of the PAtch-COvariance algorithm (PACO; Flasseur 363 et al. 2018) in its robust angular and spectral differential imaging 364 (ASDI) mode (Flasseur et al. 2020a,b). 365

PACO is a post-processing algorithm that employs ASDI to 366 model the spatial and temporal fluctuations of the image back-367 ground inside small patches through a combination of weighted 368 multivariate (i.e., accounting for the spatio-spectral correlations 369 of the speckles field) Gaussian components. Extensive testing 370 proved that the resulting SNR map is distributed as a normal-371 ized Gaussian $\mathcal{N}(0, 1)$, hence naturally providing a statistically 372 grounded detection map upon which > 5σ detections can be 373 identified at a controlled false alarm rate (e.g., at 5σ significance 374 level). The algorithm was shown to be photometrically accurate 375 and robust to false positives, and to outperform reduction meth-376 ods that are routinely employed for SPHERE (Chomez et al. 377 2023). For these reasons, PACO was appointed by the SHINE 378 consortium as the main reduction algorithm for the final analysis 379 of the whole survey (Chomez et al. 2024). 380

In addition to this, the ASDI mode of PACO uses vectors 381 of spectral weights (hereafter spectral priors) to maximize the 382 383 detection capability of candidate sources exhibiting physically 384 representative substellar spectra. As detailed in Chomez et al. 385 (2023), for every star we generated 20 such priors starting from exoplanet spectra from the ExoREM library (Charnay et al. 386 2019) ($T_{\text{eff}} \in [400, 2000]$ K) and suitable stellar spectra from 387 the BT-Nextgen AGSS2009 library (Allard et al. 2011). 388

As in Chomez et al. (2023), extensive injection tests were 389 390 performed on GPI datasets in order to ensure the reliability of 5σ detection limits, an output provided by PACO after the reduction 391 (Flasseur et al. 2020a). After randomly picking a sample of 10 392 sequences, 12 synthetic sources were evenly injected in each ob-393 servation's FOV; the mean flux of each source was set equal to 394 the 5σ detection limits estimated by PACO at the corresponding 395 coordinates. The whole process was repeated three times, vary-396 ing the input spectrum - a flat contrast spectrum, a T-type spec-397 trum and an L-type spectrum9 - of injected sources, yielding a 398 total 360 injected sources. The median SNR of the recovered 399 sources is 5.1, with little variation with spectral type, confirming 400 the statistical reliability of the contrast and detection confidence 401 estimated by PACO and underlying the statistical analysis. 402

Figure 2 shows the final performances attained by the PACO reduction as 5σ detection limits. It is possible to notice that the usage of physically-motivated spectral priors does indeed en-



Fig. 2: 5σ detection limits obtained with PACO. Individual curves are plotted in gray. The median curve is plotted as a light blue solid line, the dashed lines representing the 16% and 84% percentiles of the curve distribution. The orange and green solid lines indicate the median detection limits assuming a T-type and an L-type spectral prior, respectively. The gray box marks the inner working angle of the coronagraphic mask.

hance detection capabilities. However, in order to easily allow 406 for comparisons with reductions performed using different algorithms, we conservatively adopt in the following analysis a flat 408 spectral prior, that is a combination of spectral channels assuming that any source has the same spectral energy distribution as 410 its star; this is equivalent to standard SDI-based algorithms. 411

2.4.3. Post processing – cADI 412

One of the underlying assumptions behind PACO is that the 413 spatial and temporal fluctuations of noise inside patches are 414 much stronger than the additional contribution given by physi-415 cal sources happening to cross the patch itself during the expo-416 sure (Flasseur et al. 2018). The assumption breaks down when 417 a bright source, such as a stellar companion, is present. In other 418 words, the algorithm is optimized to detect faint sources but can 419 severely subtract, or even cancel, very bright sources in the de-420 rived SNR map. In order to complete the census of sources at 421 the bright end, we developed a custom routine based on classi-422 cal angular differential imaging (cADI; Marois et al. 2006) and 423 performed a uniform reduction of the archive. After computing 424 the pixel-wise median frame of the exposure sequence, the rou-425 tine subtracts it from every frame, then de-rotates the frames and 426 sums them up both temporally and along wavelength. The 4D 427 PSF is stacked along the temporal axis to build a 3D (x, y, λ) 428 PSF, which is then fitted by a 2D Gaussian model. The reduced 429 map is finally normalized by the peak of the PSF model so as 430 to translate it in contrast units. Detections are automatically per-431 formed on the derived map by computing the variance across 432 annuli, centered on the target star, of width equal to 1 px, and 433 finding the pixels of the map beyond a certain threshold level 434 κ (expressed in noise standard-deviation units). In a subsequent 435 step, a more precise characterization by fitting the PSF model 436 provides the astrometry and photometry of each source. 437

Because of the simplicity of the noise-reduction approach, 438 the SNR distribution in any annulus of a given contrast maps usually shows large deviations from Gaussianity. On the one 440

⁸ https://sphere.osug.fr/spip.php?rubrique16&lang=en

⁹ Input contrast spectra, created using the same procedure as spectral priors, are arbitrarily defined using the following parameters: $(T_{\text{eff}}, \log g, Z/Z_{\odot}, \text{CO}) = (1000 \text{ K}, 4.0, 1.0, 0.5)$ for the T type and $(T_{\text{eff}}, \log g, Z/Z_{\odot}, \text{CO}) = (1900 \text{ K}, 3.0, 10.0, 0.6)$ for the L type.



Fig. 3: Comparison between PACO and cADI performances. Sources only detected by PACO are shown as blue circles, while sources only seen through cADI are indicated as red diamonds. Common sources are plotted as blue squares with a red edge.

side, this issue implies that high thresholds ($\kappa \ge 20$) had to be adopted to ensure the stability of the detection step; on the other side, the poor robustness against outliers intrinsically prevents one from precisely defining a statistically grounded detection threshold.

Visual inspection of all the maps ensured the reliability of
the detections; given the above-mentioned caveats and the neglect of a correction for signal self-subtraction, the derived photometry will only be used to characterize the stellar companions
presented in Section D.

451 3. Results

452 3.1. Exoplanet candidates

The PACO reduction of our sample yielded 91 detected sources. 453 This number does not include a few false positives that could 454 be recognized and removed (see Section 3.2). 11 additional 455 sources were detected through cADI. 62 sources are detected by 456 both methods, ensuring the overlap of the respective dynamical 457 ranges. Figure 3 shows the sources detected by the two meth-458 ods. Astrometric and photometric details for all the candidates 459 460 are provided in Table C.1.

Candidate companions in DI observations are always seen as 461 unresolved point-like sources, and no information on their dis-462 tance can be discerned from a single observation; in other words, 463 it is not clear a priori if a source is physically bound to the tar-464 get star or is instead a distant background star that happens to 465 be projected close to the target. If two or more epochs are avail-466 able, the differential motion between the foreground target star 467 (and the objects bound to it) and faraway background stars can 468 be disentangled (Figure 4). 469

Whenever more than one observation was available in our sample, or if additional epochs from SPHERE could be recovered, it was possible to ascertain the proper motion of the candidates: 57 sources from the PACO reduction and 2 sources only detected with cADI were ruled out as background contaminants in this way.

If only a single observation was available, or if detection limits allowed for detection in just one epoch, we adopted an alter-



Fig. 4: Example of proper motion diagram. The astrometric displacement of the candidate around HD 84330B between the first and the second epoch is compatible with a background source with null motion (empty star). A bound object would have been in a position close to that marked by the filled star and within the boundaries allowed by a Keplerian motion.

native vetting criterion that exploits color-magnitude diagrams 478 (CMDs) to identify sources showing similar colors to known 479 imaged planets and to set them apart from background sources. 480 It might be argued that, given the availability of contrast spec-481 tra, a spectrum-based classification could be employed: how-482 ever, we argue that such a method would be highly sensitive to 483 both random uncertainties and the ignorance about the amount 484 of interstellar extinction to be adopted for background-star spec-485 tra. Conversely, the photometric method based on CMDs has 486 already been shown to be highly reliable for absolute magni-487 tudes $H \gtrsim 15$ mag, using an unprecedented sample of ~ 2000 of 488 confirmed astrophysical background sources found in SPHERE 489 data (Chomez et al. 2024). This usage of the CMD has been in-490 troduced by the SHINE consortium (Chauvin et al. 2017) and 491 its construction is fully detailed in Bonnefoy et al. (2018). This 492 tool has already been used to efficiently classify some of the 493 sources detected in the first part of the SHINE survey (Langlois 494 et al. 2021). As a first step, the H-band spectrum of each tar-495 get star was estimated by means of synthetic stellar spectra from 496 the BT-Nextgen AGSS2009 library¹⁰ (Allard et al. 2011), ade-497 quately degraded to match the spectral resolution of GPI. The 498 best-matching synthetic spectrum was identified as the closest 499 in effective temperature; the latter was empirically estimated as 500 the median value across all literature measurements found in 501 VizieR (Ochsenbein et al. 2000). Contrast spectra from candi-502 date sources detected with PACO¹¹ could thus be turned into 503 physical spectra by multiplying them by their corresponding pri-504 maries spectra. We convolved these spectra with SPHERE H2 505 and H3 filters to derive synthetic H2 and H3 photometry for 506 all our candidates; in other words, GPI spectroscopy was turned 507 into SPHERE-like photometry both to exploit the CMD vetting 508

¹⁰ The library, available at http://svo2.cab.inta-csic.es/ theory/newov2/index.php?models=bt-nextgen-agss2009, is defined by the following astrophysical parameters: $\log g[\text{cm s}^{-2}] = 4.5$, $\log Z/Z_{\odot} = 0$, alpha enhancement = 0.

¹¹ The status of all cADI candidates but one could be confirmed through dynamical arguments; the remaining one is too bright to allow for the CMD test.



Fig. 5: CMD of the companion candidates detected in this work. Overplotted to known substellar objects (white squares), background stars are represented as yellow stars if identified through proper motion analysis, or as blue circles if recognized via their color. Ambiguous sources are marked as red crosses. The exclusion area (gray) is defined by the two dashed lines. The two promising candidates (from left to right: C1 (HD 24072), C2 (HIP 78663)) are indicated as red dots.

method and to enable future comparisons between the results
from the two instruments. The convolution was possible thanks
to the broad extent of GPI's H band, whose wavelength window
covers both SPHERE narrow-band H filters.

In this way, it was possible to place every PACO candidate 513 in a (H2-H3, H2) CMD (Figure 5). We used confirmed back-514 ground objects from the SHINE survey - that offers a larger 515 sample statistics thanks to the wide 11"x11" field of view of 516 IRDIS (Dohlen et al. 2008) – to build an "exclusion zone", de-517 fined as the region of the CMD that encompasses all the points 518 within 5σ from the mean colors of background sources as a func-519 tion of their absolute magnitude. As in SHINE publications, the 520 exclusion zone was set to begin at H2 = 16 mag, as the exis-521 tence of some planets (e.g., HR 8799 b) with $H2 \sim 15$ mag and 522 $H2-H3 \sim 0$ mag renders the method unreliable at brighter mag-523 nitudes (Langlois et al. 2021, Chomez et al. 2024). We labeled 524 as "companion candidates" the sources lying along the T track 525 or having additional indications hinting towards a bound nature, 526 and as "ambiguous" the sources in the region H2 < 16 mag and 527 $H2 - H3 \sim 0.$ 528

529 Excluding already known substellar companions, all but nine 530 sources can be confidently ruled out as background contami-531 nants. Seven of these are classified as ambiguous according to

our vetting scheme, and will not be hereafter discussed. The na-532 ture of the two remaining promising candidates - whose photom-533 etry and age is consistent with 5-8 M_{Jup} objects – is currently 534 unclear. The candidate around HIP 78663 is located in a posi-535 tion of the CMD where the colors of bound companions overlay 536 those of background stars; however, we classify it as a promis-537 ing candidate because of a tentative ~ 3.5σ detection in the shal-538 lower second epoch possibly hinting at common proper motion. 539 As regards the candidate around HD 24072, in addition to the 540 hypothesis of a bound nature, the following scenarios might be 541 envisaged to explain its position along the young-object track: 542

- a free-floating planet or brown dwarf, belonging to the same association as the target and hence possessing similar colors to substellar companion while not exhibiting a large variation of the distance modulus;
- a statistical false positive (see Section 3.2). A spectral dependence of the photocenter of a false positive might happen to mimic, during the characterization step of PACO, blue spectra similar to those of real substellar companions lying along the T track.

Spurious detections in direct imaging have previously arisen due to extended objects (proto-planetary and debris disks) that were poorly subtracted (see, e.g., Sallum et al. 2015 and confutation by Currie et al. 2019), but we exclude this possibility given the lack of infrared excess in WISE (Wright et al. 2010) bands.

We finally notice that the HD 24072 system also comprises 557 a low-mass star, closer to the primary than the planet candidate 558 (see Section D); under the assumption of face-on circular orbits, 559 we empirically verified, based on the results by Musielak et al. 560 (2005), that the candidate would be far enough from the substellar companion to be dynamically stable. 562

Our reanalysis redetected all substellar companions (7 plan-563 ets, 3 brown dwarfs) that we expected to find on the basis of 564 the literature (Figure 6). Some of these companions - notably, 565 HR 8799 c, d, and e – have just one epoch in our observing 566 sample; consistently with the decision tree described above, we 567 would have been able to confirm them as bound objects through 568 proper motion test, employing additional available SPHERE or 569 GPI epochs. 570

We report in Table 1 details about the astrometry and the photometry of these 10 substellar companions. Given the extensive characterization of these objects already undertaken in the literature, we deem a re-derivation of masses and semi-major axes – the main input needed for the statistical analysis of Section 3.4 - to be outside of the scope of this paper; instead, we decided to recover the most accurate values from dedicated literature works. 577

In addition to substellar companions, we detected 6 sources 578 whose high luminosity points towards a stellar nature. We were 579 able to confirm 5 of them as physically bound thanks to 1) a 580 proper motion strongly disagreeing with background stars, and 581 2) astrometric wobbles indicated by the Gaia astrometric solu-582 tion of the corresponding primaries; the remaining binary can-583 didate to HD 74341B, with no second epoch and too bright to 584 employ the CMD test, awaits confirmation. Details are provided 585 in Section **D**. 586

3.2. False positives

As mentioned in Section 3.1, the roster of PACO candidates does not include a few detections identified as false positives, induced either 1) by real astrophysical or optical features, or 2) by statistical fluctuations of the SNR map. The former category includes residuals of the first Airy ring around very bright sources 592

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 Δx [mas]

Fig. 6: PACO or cADI detection maps for the substellar companions detected in the survey (indicated by arrows). PACO maps are to be read as SNR maps, sharing a common colorbar. Individual colorbars are shown below the two cADI maps.

Table 1: Detected substellar companions to stars in the sample.

Name	subsample	Date	SNR	sep	PA	ΔH	Н	H2 – H3	mass	sma	source
	-			arcsec	deg	mag	mag	mag	$M_{\rm Jup}$	au	
HD 206893 b	FGK	2016-09-22	7.2	0.268 ± 0.002	61.5 ± 0.4	11.1	13.7 ± 0.1	0.5 ± 0.1	28	9.6	1,1
HR 8799 c	BA	2016-09-19	37.1	0.953 ± 0.002	330.8 ± 0.2	12.5	14.8 ± 0.2	0.4 ± 0.2	8.5	41	2,2
HR 8799 d	BA	2016-09-19	19.4	0.666 ± 0.002	223.5 ± 0.2	12.5	14.7 ± 0.2	0.3 ± 0.2	8.5	27	2,2
HR 8799 e	BA	2016-09-19	9.2	0.393 ± 0.003	284.1 ± 0.4	12.2	14.4 ± 0.2	0.3 ± 0.2	9.6	16	3,2
		2014-12-18	14.7	0.439 ± 0.002	171.3 ± 0.3	14.3	16.7 ± 0.2	-1.2 ± 0.1			
		2015-01-31	5.0	0.456 ± 0.007	170.4 ± 0.8	15.1	_	—			
51 Eri b	BA	2015-09-01	10.9	0.442 ± 0.003	166.8 ± 0.4	14.7	17.1 ± 0.2	-1.0 ± 0.2	4.1	11.1	4,5
		2015-12-20	7.5	0.443 ± 0.005	166.5 ± 0.7	14.4	16.8 ± 0.2	-0.6 ± 0.2			
		2016-09-18	12.7	0.442 ± 0.003	162.0 ± 0.4	14.4	16.8 ± 0.1	-0.8 ± 0.1			
		2015-11-06	8.3	0.421 ± 0.004	359.1 ± 0.6	9.8	11.9 ± 0.2	0.4 ± 0.1			
β Pic b	BA	2015-12-22	7.7	0.241 ± 0.002	213.2 ± 0.4	9.6	11.7 ± 0.2	0.4 ± 0.1	11.9	9.93	6,6
		2016-01-21	7.5	0.226 ± 0.002	212.3 ± 0.6	10.4	12.5 ± 0.3	0.4 ± 0.3			
		2013-12-11	8.9	0.635 ± 0.003	150.6 ± 0.3	13.8	15.9 ± 0.2	0.2 ± 0.2			
UD 05086 b	ВV	2016-02-29	9.3	0.627 ± 0.003	148.1 ± 0.4	13.6	15.8 ± 0.2	0.2 ± 0.1	26	617	78
11D 95080 0	DA	2016-03-06	8.1	0.626 ± 0.004	148.1 ± 0.4	14.0	16.1 ± 0.2	0.7 ± 0.2	2.0	01.7	7,0
		2016-03-28	4.9	0.629 ± 0.006	147.5 ± 0.6	14.2	16.4 ± 0.2	0.0 ± 0.3			
		2016-01-25	17.0	0.605 ± 0.002	294.7 ± 0.3	12.8	15.3 ± 0.2	0.6 ± 0.3			
UD 2562 h	FGK	2017-02-13	21.4	0.635 ± 0.002	298.3 ± 0.2	11.3	13.8 ± 0.1	0.3 ± 0.1	10.28	21.2	0.0
TIK 2502 0	TOK	2017-11-29	12.5	0.654 ± 0.002	297.8 ± 0.3	12.8	15.2 ± 0.2	0.4 ± 0.1	10.28	21.2	9,9
		2018-11-19	28.1	0.677 ± 0.002	297.3 ± 0.2	11.3	13.7 ± 0.3	0.3 ± 0.3			
HD 984 B ^a	FGK	2015-08-30	_	0.219 ± 0.002	84.0 ± 0.4	7.7			61	28	10,10
PZ Tel B ^a	FGK	2015-07-30		0.501 ± 0.002	$5\overline{9.5 \pm 0.2}$	5.9			27	27	11,11

Notes. ^{*a*}: reduction through the custom cADI. (mass, sma) sources for planet properties: 1: Hinkley et al. (2023); 2: Zurlo et al. (2022); 3: Brandt et al. (2021); 4: Elliott et al. (2024); 5: De Rosa et al. (2020b); 6: Lacour et al. (2021); 7: Nielsen et al. (2019); 8: Rameau et al. (2016); 9: Zhang et al. (2023); 10: Franson et al. (2022); 11: Franson & Bowler (2023). The column named "subsample" indicates whether the parent star belongs to the FGK or the BA subsample (see Sec. 3.4).

(2 cases) and disk residuals (10 cases); the latter (8 cases) is con-stituted by unusually bright residuals that had no counterpart in

additional GPI or SPHERE observations with better or similar 595 detection limits. 596

With respect to the latter case, we tried to estimate the num-597 ber of false positives expected to arise from statistical fluc-598 tuations. We recall that the distribution of pixel intensities in 599 PACO SNR maps is a normalized Gaussian $\mathcal{N}(0, 1)$. In this 600 case, a 5σ threshold corresponds to a false alarm probability 601 $p_{5\sigma} = 2.9 \cdot 10^{-7}$. Given the number of pixels in GPI's FOV, 602 $N_{px} = 185^2$ and the number of effectively independent spectral 603 priors, N_p , that we empirically estimate as $N_p \approx 4^{12}$, using a 604 binomial distribution, we expect ~ 20 false positives across the entire survey (see Chomez et al. 2023). This number is larger 605 606 than the number of statistical false positives that could be iden-607 tified through second-epoch observations; hence, we expect that 608 some sources labeled as CMD background sources also belong 609 to the category. 610

611 3.3. Completeness

The completeness of our survey was quantified in the following way. As a first step, we azimuthally averaged the 2D detection maps provided by PACO, obtaining 1D contrast curves (detection limits at 5σ).

Pending a final confirmation of the nature of the two promis-616 ing candidate companions, the detection limits of the corre-617 sponding observations were adjusted accordingly to ensure the 618 statistic reliability of the corresponding observations. The same 619 was done for the seven datasets containing ambiguous sources. 620 In particular, the mean contrast of each candidate was employed 621 as a floor value in the corresponding 2D 5 σ map. In other words, 622 we pretend to have shallower observations so that a source as 623 bright as the candidate can be at most a marginal 5σ detection. 624 These maps were then collapsed to 1D as described above. 625

The 1D curves obtained in this way were converted into 626 mass limits through MADYS, adopting the stellar parameters dis-627 cussed in Section 2.3. The mass-luminosity relation is based on 628 the ATMO evolutionary models (hereafter ATMO; Phillips et al. 629 2020; Chabrier et al. 2023)¹³. Chemical disequilibrium is ex-630 pected to critically affect the atmospheric features of cool T-type 631 and Y-type objects (see, e.g. Leggett et al. 2015, 2017; Miles 632 633 et al. 2020; Baxter et al. 2021); given that 1) the corresponding 634 temperature range is within the reach of our analysis (see Fig-635 ure 5), and 2) the effect is particularly strong in H-band observa-636 tions such as those under consideration, we decided to employ the grid assuming weak chemical disequilibrium (ATMO-NEQ-637 W) instead of chemical equilibrium (ATMO-CEQ). We explored 638 in Appendix E the effect of this assumption, comparing the re-639 sults with those obtained under chemical equilibrium and strong 640 disequilibrium (ATMO-NEO-S)¹⁴. In addition to this, the impact 641 of model selection and age uncertainty was quantified. 642

Starting from mass limits, the completeness could be estimated through Exo-DMC¹⁵ (Bonavita 2020). Within each cell of a 2D grid in the (mass, sma) plane, the detectability of N = 1000companions to every star, whose orbital parameters are drawn in a Monte Carlo fashion, is computed by comparison with 5σ

(0) $(10^{10}$ (10^{10}) $(10^{10}$ (10^{10}) $(10^{$

Fig. 7: Survey completeness as a function of companion mass and semi-major axis, computed using the ATMO-NEQ-W models. Red stars indicate known substellar companions (see Table 1).

Table 2: Input parameters used for Exo-DMC.

parameter	description
no. of steps in sma	500
no. of steps in mass	200
no. of draws per cell	1000
semi-major axis	log-uniform in [0.1, 1000] au
companion mass	log-uniform in $[0.1, 100] M_{Jup}$
inclination	$\cos i$ uniform in $[-1, 1]$
eccentricity	N(0, 0.9)
longitude of node	uniform in $[0, 2\pi]$
longitude of periastron	uniform in $[0, 2\pi]$
fraction of period	uniform in [0, 1]

mass limits as a function of the projected separation. We provide 648 in Table 2 an overview of the adopted parameters. 649

The final map is computed by taking the average of all the 650 individual maps. When multiple epochs for a given star are avail-651 able, the largest value for the detectability is selected for every 652 cell. The results are shown in Figure 7. We notice that the peak 653 sensitivity of the survey is about 88%: we interpret such a low 654 value as the combination of three factors: 1) the small field of 655 view of the instrument; 2) the moderate distance spread across 656 the sample; 3) the fact that, working in semi-major axis and not 657 in projected separation, a fraction of planets with given a might 658 be undetectable because of projection effects. 659

We are now able to directly compare our detection capabil-660 ities with those of N19, so as to justify a posteriori the idea of 661 a reanalysis of those archival data. In order to avoid any possi-662 ble systematic difference, a new map was computed only using 663 the observations considered therein; moreover, instead of using 664 Exo-DMC, we closely reproduced the original method, includ-665 ing the same values for distances, ages and substellar evolution-666 ary model. The comparison, shown in Figure 8, indicates that 667 the PACO-based reanalysis allows for a significant performance 668 gain at all separations, which can be up to twofold in terms of 669 detectable mass at given completeness. 670

¹² The correlation between SNR maps under any two spectral priors is larger than zero. By "number of effectively independent spectral priors" we mean the ratio $N_{positives, 20 priors}/N_{positives, 1 prior}$, estimated through extensive testing.

¹³ Using the most recent version, that features a new equation of state for dense hydrogen-helium mixtures: https://noctis.erc-atmo.eu/fsdownload/zyU96xA6o/phillips2020.

¹⁴ The amount of vertical mixing in disequilibrium models is parametrized through the eddy diffusion coefficient K_{ZZ} . Constraining K_{ZZ} is a long-standing issue (see discussion in Phillips et al. 2020). ¹⁵ https://github.com/mbonav/Exo_DMC



Fig. 8: Comparison of the survey completeness between N19 (dashed lines) and this work (solid lines). Only the observations used by N19 were used to draw this plot.

671 3.4. Planet occurrence rates

Deriving unbiased occurrence frequencies of exoplanets is one of the main goals of large blind surveys, and in turn a crucial input to draw comparisons with formation models. Provided a large enough sample, it is additionally possible to investigate the dependence of these frequencies on host properties such as mass and metallicity, highlighting the key role of the parent star in shaping its planetary system.

We begin our investigation by focusing on the occurrence frequency f for the entire stellar population represented by the GPIES sample. Extracting this quantity from the fact of having observed N companions given a certain survey completeness is a typical inversion problem that can be treated within a Bayesian framework.

We employ a formalism that is similar to that used in previ-685 ous direct imaging studies (see, e.g. Lafrenière et al. 2007; Lan-686 nier et al. 2016). Given a certain area \mathcal{A} in the (sma, mass) plane 687 defined by $a_{\min} < a < a_{\max}$ and $m_{\min} < m < m_{\max}$, let us de-688 fine p_i the mean probability to see a companion around the *i*-th 689 star lying within \mathcal{A} . Based on our completeness analysis (Sec-690 tion 3.3), p_i can be estimated as the mean detection probability 691 in \mathcal{A} across the entire survey, that is the mean value in \mathcal{A} of the 692 completeness map shown in Figure 7. 693

The probability $p_{det,i}$ to detect a companion in \mathcal{A} around the *i*-th star is the product of the detection probability and the underlying occurrence frequency $f: p_{det,i} = p_i \cdot f$. The connection with the observed planet sample is mediated by d, a vector whose *i* element represents the number of companions detected within \mathcal{A} around the *i*-th star.

The likelihood of the observed data as a function of the f can be estimated as the product of individual Bernoulli events, one per star:

$$L(\{d_i\}|f) = \prod_{i=1}^{N} (1 - p_{\det,i})^{1 - d_i} \times (p_{\det,i})^{d_i}$$
(2)

The probability density function of f, that is the occurrence703frequency of companions in \mathcal{A} given the data can be finally esti-704mated through Bayes' theorem:705

$$P(f|\{d_i\}) = \frac{L(\{d_i\}|f)P(f)}{\int_0^1 L(\{d_i\}|f)P(f)df}$$
(3)

as the posterior distribution emerging from the interplay 706 between a suitable prior distribution P(f) and the likelihood 707 $L(\{d_i\}|f)$. We adopt two distinct priors: a uniform prior and a 708 Jeffreys prior. The uniform prior: 709

$$P(f) \propto 1, \qquad \qquad \forall f \in [0,1]$$
 (4)

despite not incorporating any observational information, is 710 not uninformative, as it assumes much larger weights for large 711 values of f compared to what is expected from observations. 712 Nevertheless, the simplicity of this prior makes it widely adopted 713 in the literature: we decided to employ it in order to allow for 714 comparison with published results. 715

A Jeffreys prior has the twofold advantage of being noninformative and counterbalancing the bias that favors $f \sim 0.5$. 717 In the case of Bernoulli events, the Jeffreys prior for the parameter f is simply: 719

$$P(f) = \frac{1}{\sqrt{f \cdot (1-f)}} \tag{5}$$

We adopt the latter prior distribution, that has the advantage 720 of being non-informative, as our standard choice in the following 721 analysis. 722

A particularly delicate point is represented by the choice of 723 \mathcal{A} : on the one hand, selecting a too narrow range would result in 724 a critical amplification of fluctuations from small number statis-725 tics; on the other hand, including regions where $\langle p_i \rangle \sim 0$ 726 would require a significant amount of extrapolation due to the 727 lack of data and, consequently, induce a flattening of the poste-728 rior distribution over the prior. An additional factor to take into 729 account is the dependence of the results on both age uncertainty 730 and model selection, becoming more severe as the lower mass 731 limit is decreased (Appendix E). We decided to consider, as our 732 nominal case, a lower mass limit of $5M_{Jup}$ and a semi-major axis 733 range 10 au < a < 100 au as a compromise between these 734 concurrent factors; the upper mass limit will be set to either 735 13 M_{Jup} or 80 M_{Jup} depending on whether brown dwarf com-736 panions are considered or not. We derive occurrence frequencies 737 of $1.7^{+0.9}_{-0.7}\%$ when $\mathcal{A} = [5, 13] M_{Jup} \times [10, 100]$ au, and $2.2^{+1.0}_{-0.8}\%$ 738 when $\mathcal{\tilde{H}} = [5, 80] M_{Jup} \times [10, 100]$ au, with f represented by 739 the median of the posterior and the error bars defined from the 740 [16%, 84%] percentiles. In order to allow a straightforward com-741 parison with previous results from the literature (Section 4), we 742 also present additional occurrences starting from different defi-743 nitions of \mathcal{A} (Table 3). We notice that no significant difference is 744 derived from the prior choice, confirming that our careful choice 745 for \mathcal{A} did minimize the impact of the prior; the errorbars, as 746 expected, are smaller in Jeffreys case compared to the uniform 747 case. In addition to this, no significant deviation arises for this 748

Table 3: Occurrence rates for different definitions of the (mass, sma) range \mathcal{A} and for the two choices for the prior distribution (*U*: uniform; *J*: Jeffreys). ^{*a*}: 95% upper limit.

${\mathcal A}$	SpT	f_U	f_J
$M_{\rm Jup} imes$ au		%	%
$[5, 13] \times [10, 100]$	all	$1.9^{+1.0}_{-0.7}$	$1.7^{+0.9}_{-0.7}$
$[5, 13] \times [10, 100]$	BA	$4.3^{+2.6}_{-1.9}$	$3.8^{+2.4}_{-1.7}$
$[5, 13] \times [10, 100]$	FGK	$1.0^{+1.0}_{-0.6}$	$0.7^{+0.9}_{-0.5}$
$[5, 80] \times [10, 100]$	all	$2.4^{+1.0}_{-0.8}$	$2.2^{+1.0}_{-0.8}$
$[5, 80] \times [10, 100]$	BA	$3.5^{+2.1}_{-1.5}$	$3.0^{+2.0}_{-1.4}$
$[5, 80] \times [10, 100]$	FGK	$2.2^{+1.3}_{-0.9}$	$1.9^{+1.2}_{-0.9}$
$[2, 13] \times [10, 100]$	all	$3.5^{+1.5}_{-1.2}$	$3.2^{+1.5}_{-1.1}$
$[2, 13] \times [3, 100]$	all	$5.3^{+2.1}_{-1.7}$	$5.0^{+2.1}_{-1.6}$
$[2, 13] \times [5, 300]$	all	$5.3^{+2.1}_{-1.7}$	$5.0^{+2.1}_{-1.6}$
$[13, 80] \times [5, 100]$	all	$1.4^{+0.9}_{-0.6}$	$1.2^{+0.8}_{-0.6}$
$[13, 80] \times [5, 100]$	BA	< 3.8 ^{<i>a</i>}	$< 2.7^{a}$
$[13, 80] \times [5, 100]$	FGK	$2.3^{+1.4}_{-1.0}$	$2.0^{+1.3}_{-0.9}$

choice of \mathcal{A} as an effect of the theoretical assumptions and observational uncertainties, ensuring the robustness of our results (Appendix E).

We consider host star metallicity not to be a factor of par-752 ticular concern, as the metallicity of young star-forming regions 753 in the solar neighborhood is typically solar with limited spread 754 (D'Orazi et al. 2011; Biazzo et al. 2012; Baratella et al. 2020; 755 Magrini et al. 2023). Conversely, as done in Nielsen et al. (2019) 756 and Vigan et al. (2021), we explicitly investigate the dependence 757 of the occurrence frequency on stellar mass. We divided our sam-758 ple in three bins of stellar masses, obtaining the BA subsample 759 $(M > 1.5M_{\odot}, 160 \text{ stars})$, the FGK subsample $(0.5 < M \le 1.5M_{\odot}, M \le 1.5M_{\odot})$ 760 235 stars), and the M subsample ($M \le 0.5 M_{\odot}$, 5 stars). Given its 761 small size, the M star sample was discarded. 762

Both the aggregated results and the mass-dependent ones are plotted in Fig 9. Occurrence frequencies for different values of \mathcal{A} are provided for reference in Table 3; moreover, a digitized version of completeness maps is also made available¹⁶ so as to allow interested readers to extract additional results based on different definitions of \mathcal{A} .

769 4. Discussion

Thanks to the analysis performed in Section 3.4, it is now possible to directly compare the results emerging from our PACO
re-reduction of GPI data to previous literature works.

Figure 10 presents a juxtaposition of our results with the 773 frequencies derived from the first 300 stars of GPIES (Nielsen 774 et al. 2019) and the first 150 stars of SHINE (SHINE F150; Vi-775 gan et al. 2021). Moreover, the results emerging from the meta-776 analysis of 384 imaged stars by Bowler (2016) are shown, al-777 though we stress that they are, by design, less protected against 778 selection biases due to the heterogeneous underlying sample. A 779 general finding is that our results are fully compatible with liter-780 ature estimates, but they are typically more precise. In particular, 781 782 the comparison with N19 clearly indicates that the new analysis 783 places the tightest constraints to date on giant planet occurrence 784 based on GPI data, with a gain in precision being a direct consequence of the large gain in completeness (Figure 8). The smaller 785



Fig. 9: Occurrence frequency of GP (upper panel) and GP+BD (lower panel) from the re-analysis of the 400-star GPIES sample presented in this work. Aggregated results are shown in blue, whereas results for the BA and the FGK subsample are plotted in green and orange, respectively. The colored area encompasses the $[16^{th}, 84^{th}]$ of the posterior distribution.

frequency of companions is an effect of the increased complete-786 ness with no new confirmed detection. As regards SHINE F150, 787 the much larger field of view of IRDIS (11"x11", compared to 788 the 2.7"x2.7" FOV of GPI) ensures a much more complete cov-789 erage of the semi-major axis range of interest and thus larger 790 room for planet detection, compensating the twofold advantage 791 of our study in terms of sample size and reduction performances: 792 as a result, the precision of the derived occurrence rates is sim-793 ilar. In this regard, the full analysis of SHINE data with PACO 794 (Chomez et al. 2024), which combines all the advantages of the 795 two analyses, is expected to provide an invaluable contribution 796 to demographic studies of wide-orbit exoplanets. 797

In view of the profound consequences with respect to planet 798 formation scenarios, it is extremely interesting to assess the 799 dependence of the observed occurrence rates on stellar mass. 800 As in N19, we employed a threshold value of $M = 1.5 M_{\odot}$ 801 to distinguish a BA subsample and a FGK subsample. That 802 study claimed that a significant (3.4σ) tension between BA 803 and FGK planet rates (f_{BA} and f_{FGK} , respectively) exists for 804 $\mathcal{A} = [2, 13] M_{Jup} \times [3, 100]$ au := \mathcal{A}_1 , with planets being more 805

¹⁶ url of Zenodo repository, to be created upon acceptance



Fig. 10: Comparison of the occurrence rates (left panel: $\mathcal{A} = [2, 13] M_{Jup} \times [3, 100]$ au; right panel: $\mathcal{A} = [5, 13] M_{Jup} \times [10, 100]$ au) of giant planets with previous analyses by Bowler (2016), Nielsen et al. (2019) and Vigan et al. (2021). The left and the right half of each panel are relative to BA and FGK stars, respectively. Estimates indicated by arrows are to be read as 95% upper limits, while error bars on point estimates are defined as to encompass the 68% C.I.. Question marks indicate missing data points.

806 common around BA hosts; while six companions were detected 807 inside \mathcal{A}_1 in the BA subsample, no companion was identified in 808 the FGK subsample. However, we expect the result to be weakened by the re-revaluation of the mass of HR 2562 b (Zhang 809 et al. 2023), a companion to an FGK star now firmly placed into 810 the planetary-mass domain. In order to verify whether this is the 811 case, we drew, in a Monte Carlo fashion, values from the poste-812 rior distributions of f_{BA} and f_{FGK} under \mathcal{A}_1 . The values from the 813 latter distribution are larger than those drawn from the former 814 in 0.1% of the cases, implying a 3.3 σ tension between the two 815 distributions. Hence, our analysis confirms the finding by N19. 816

With respect to brown dwarf companions, no statistically sig-817 nificant difference in the observed rates was found by N19 be-818 tween BA and FGK hosts. The observation is in line with the 819 results of previous analyses showing compatible rates across a 820 wide range of stella types (see, e.g., Nielsen et al. 2013; Bowler 821 et al. 2015; Lannier et al. 2016; Bowler & Nielsen 2018). Based 822 on our analysis, a tentative (1.7σ) tension between the two rates 823 is found for $\mathcal{A} = [13, 80] M_{Jup} \times [5, 100]$ au, with an interest-824 ing inversion compared to the planetary case: in other words, 825 giant planets appear to be more common around BA host, while 826 brown dwarf companions tentatively appear to be more common 827 around FGK hosts. 828

Although the BD trend is not statistically significant, an in-829 teresting analogy might be drawn with the behavior of the two 830 empirical distributions of substellar companions introduced by 831 Vigan et al. (2021) in the context of SHINE. A planet-like and 832 a star-like distribution of companions – both being the product 833 of a log-normal distribution for semimajor axis and a power-law 834 for companion-to-star mass ratios - were simultaneously fitted 835 to the substellar companion population, divided in three bins of 836 mass (BA, FGK, M). Similarly to our planetary rates, the me-837 dian values of the planet-like posterior are larger than those of 838 the star-like posterior for BA hosts, and smaller for FGK hosts. 839 It might be argued that a strict distinction between giant planets 840 and brown dwarfs based on the deuterium burning limit is not 841 adequate to capture the complexity of the different formation 842 mechanisms (CA and GI) involved (Chabrier et al. 2014), and 843

that studying together the entire population is the key to identify population trends (see, e.g., Gratton et al. 2024); while this is certainly true, first-order, population-wise differences in some parameter, arguing for different underlying formation channels, can sometimes be discerned using rough mass boundaries (Bowler et al. 2020).

In view of the low occurrence rates, the large extent of host 850 star masses, the interplay of different formation channels and the 851 impact of input assumptions, we deem it necessary to defer a 852 thorough study of the distribution of companion properties to our 853 future joint SPHERE+GPI analysis: thanks to its larger sample 854 size, this sample is expected to bring about much tighter con-855 straints on the properties of the companion population, offering 856 in turn the possibility to compare them both to empirical distri-857 butions and to synthetic populations of companions produced by 858 formation models. 859

5. Conclusions

We have presented in this work the results of a complete rereduction of 400 stars from the GPIES survey, one of the largest planet-hunting DI endeavors to date, by means of an advanced post-processing algorithm named PACO. The key results of this work are the following: 863

- the detection capabilities of the survey were greatly enhanced by means of our novel post-processing technique, reaching up to a twofold gain in terms of detectable mass at given completeness;
- out of 102 detected sources, 2 were identified as promising companion candidates awaiting follow-up confirmation;
 871
- thanks to the deeper detection limits provided by PACO, it was possible to place some of the deepest constraints ever provided by direct imaging on the occurrence of wide-orbit giant planets. We derive an occurrence rate of $1.7^{+1.0}_{-0.7}\%$ for $5 M_{Jup} < m < 13 M_{Jup}$ planets in 10 au < a < 100 au, increasing to $2.3^{+1.0}_{-0.8}\%$ when including substellar companions up to 80 M_{Jup} ;

- we verified that the above-mentioned results are robust 879 against the effect of age uncertainty, model selection, and 880 disequilibrium chemistry; 881
- as in previous studies, we observe (3.3 σ C.L.) a larger oc-882 currence rate of giant planets around BA hosts compared to 883 FGK stars; 884
- we tentatively (1.7 σ C.L.) identify an inversion of this trend 885 when considering brown dwarf companions, with FGK stars 886 possibly hosting more such companion than their BA coun-887 terparts. 888

In a forthcoming study, we plan to combine the archives of 889 SPHERE and GPI data, leading to a threefold sample size com-890 pared to this work. By applying the same reduction and analysis 891 methods presented here, it will be possible to assess a whole 892 series of stimulating questions related to the origin, the preva-893 lence and the properties of wide-orbit planets. In addition to this, 894 these endeavors will enable a decisive step towards the coveted 895 combination of demographic constraints derived through differ-896 897 ent detection techniques, delivering in turn key inputs for planet

898 formation models suited to a wide variety of host stars.

- Acknowledgements. This project has received funding from the European Re-899 900 search Council (ERC) under the European Union's Horizon 2020 research and innovation program (COBREX; grant agreement nº 885593). 901
- 902 SPHERE is an instrument designed and built by a consortium consisting of IPAG 903 (Grenoble, France), MPIA (Heidelberg, Germany), LAM (Marseille, France), LESIA (Paris, France), Laboratoire Lagrange (Nice, France), INAF - Osser-904 vatorio di Padova (Italy), Observatoire de Genève (Switzerland), ETH Zürich 905 (Switzerland), NOVA (Netherlands), ONERA (France) and ASTRON (Nether-906 907 lands) in collaboration with ESO. SPHERE was funded by ESO, with additional contributions from CNRS (France), MPIA (Germany), INAF (Italy), FINES 908 909 (Switzerland) and NOVA (Netherlands). SPHERE also received funding from 910 the European Commission Sixth and Seventh Framework Programmes as part of the Optical Infrared Coordination Network for Astronomy (OPTICON) under 911
- grant number RII3-Ct-2004-001566 for FP6 (2004-2008), grant number 226604 912 913 for FP7 (2009-2012) and grant number 312430 for FP7 (2013-2016).
- This work has made use of the High Contrast Data Centre, jointly operated by 914 915 OSUG/IPAG (Grenoble), PYTHEAS/LAM/CeSAM (Marseille), OCA/Lagrange
- (Nice), Observatoire de Paris/LESIA (Paris), and Observatoire de Lyon/CRAL, 916 917 and is supported by a grant from Labex OSUG@2020 (Investissements d'avenir 918 ANR10 LABX56).
- 919 This work is based on observations obtained at the Gemini Observatory, which is operated by the Association of Universities for Research in Astronomy, Inc., un-920 921 der a cooperative agreement with the NSF on behalf of the Gemini partnership:
- the National Science Foundation (United States), the National Research Coun-922 cil (Canada), CONICYT (Chile), the Australian Research Council (Australia), 923 Ministério Ciência, Tecnologia e Inovação (Brazil) and Ministerio de Ciencia, 924
- 925 Tecnología e Innovación Productiva (Argentina). 926 This research has made use of data obtained from or tools provided by the portal
- exoplanet.eu of The Extrasolar Planets Encyclopaedia. 927 928
- This research has made use of the VizieR catalogue access tool, CDS, Strasbourg, France (DOI: 10.26093/cds/vizier). The original description of the VizieR 929 service was published in 2000, A&AS 143, 23. 930
- This research has made use of the SIMBAD database, CDS, Strasbourg Astro-931 nomical Observatory, France. The original description of the SIMBAD database 932 was published in 2000, A&AS 143, 9. 933
- This work has made use of data from the SHINE GTO survey, operated at 934 935 SPHERE@VLT.
- This research has made use of the SVO Filter Profile Service "Carlos Ro-936 drigo", funded by MCIN/AEI/10.13039/501100011033/ through grant PID2020-937
- 938 112949GB-I00.

We thank Schuyler Grace Wolff for her help to calculate GPI calibration files 939 940 using GPI DRP.

- We are grateful to the anonymous referee for the insightful comments provided 941 942 during the peer-review, which largely contributed to raising the quality of the 943 manuscript.
- 944 Software: numpy (Harris et al. 2020), astropy (Astropy Collaboration et al. 2013), astroquery (Ginsburg et al. 2019), madys (Squicciarini & Bonavita 945 946 2022), GaiaPMEX (Kiefer et al. 2024), pyklip (Wang et al. 2015b), Exo-DMC
- 947 (Bonavita 2020), PACO (Flasseur et al. 2018).

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Article number, page 13 of 21

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Article number, page 14 of 21

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Appendix A: The stellar sample

Table A.1: Stellar properties for the sample considered in this work. The full table is available in electronic form at the CDS.

star name	ra ^a	dec^a	parallax ^a	SpT^b	Gmag^a	$Hmag^c$	E(B-V)	YMG	$p_{\rm memb}$	age	age ref. ^d	mass
	hms	dms	mas		mag	mag	mag			Myr		M_{\odot}
HD 104467	12 01 39.1168	-78 59 16.915	10.18 ± 0.12	G3V(e)	8.44	6.97	0.02	EPSC	1.00	$3.7^{+4.6}_{-1.4}$	В	$1.74^{+0.09}_{-0.24}$
HD 105874A	12 11 14.8135	-52 13 03.187	8.07 ± 0.99		7.88		0.04	LCC	0.99	15 ± 3	В	1.69 ± 0.08
HD 118991A	13 41 44.7704	-54 33 33.934	11.28 ± 0.10	B8.5Vn	5.24	5.45	0.02	FIELD	0.72	15.0 ± 3.8	Ν	3.32 ± 0.17
HD 129926B	14 46 00.5907	-25 26 39.973	32.62 ± 0.02	G1V	6.95	5.72	0.01	FIELD	0.71	500 ± 120	Ν	1.08 ± 0.05
HD 131399A	14 54 25.3089	-34 08 34.038	10.20 ± 0.70	A1V	7.07		0.04	UCL	0.99	16 ± 2	В	1.94 ± 0.10
HD 137919A	15 30 21.31	-41 55 08.33	7.93 ± 0.72			6.46	0.03	UCL	0.99	16 ± 2	В	3.54 ± 0.18
HD 141943	15 53 27.2916	-42 16 00.71	16.63 ± 0.02	G2	7.79	6.41	0.01	FIELD	0.48	16 ± 4	Ν	$1.22^{+0.09}_{-0.06}$
HD 147553A	16 23 56.7146	-33 11 57.828	7.23 ± 0.04	B9.5V(n)	7.00	7.01	0.04	UCL	0.95	16 ± 2	В	2.51 ± 0.13
HD 16699A	2 38 44.2802	-52 57 03.053	17.27 ± 0.02	F8V	7.75	6.70	0.00	ARG	0.98	45 ± 5	В	1.22 ± 0.06
HD 16699B	2 38 45.0461	-52 57 08.451	16.69 ± 0.21	G8V	8.24	6.63	0.00	ARG	0.95	45 ± 5	В	1.15 ± 0.06

Notes. Data taken from: ^{*a*}: Gaia DR3; ^{*b*}: Simbad; ^{*c*}: 2MASS; ^{*d*}: (B)ANYAN, (N)ielsen+19, (S)HINE. Details about the derivation of E(B-V), 1173 YMG membership, ages and masses are provided in Section 2.3.

Table A.2: Adopted ages for the YMG of interest.

YMG	Age	Source
AB Doradus	149^{+31}_{-49}	1
Argus	45 ± 5	2
β Pic MG	24 ± 5	1
Carina	45^{+11}_{-7}	3
Carina-Near	200 ± 50	4
Columba	42^{+6}_{-4}	3
ϵ Cha	$3.7^{+4.6}_{-1.4}$	5
Lower Centaurus-Crux	15 ± 3	6
Tucana Horologium Ass.	45+5	3
TW Hya Ass.	10 ± 3	3
Upper Centaurus-Lupus	16 ± 2	6
Upper Scorpius	10 ± 3	6
Volans-Carina Ass.	87^{+5}_{-9}	7

Notes. Sources: 1: Desidera et al. (2021); 2: Zuckerman (2019); 3: Bell et al. (2015); 4: Zuckerman et al. (2006), assuming a relative 25% error; 5: Murphy et al. (2013); 6: Pecaut & Mamajek (2016); 7: Gagné et al. (2018a).

1175 **Appendix B: Observation logs**

Table B.1: Observing log for the observations considered in this work. The full table is available in electronic form at the CDS.

star name	GPIES name	obs. night	$ au_0{}^a$	$seeing^b$	airmass	int. time ^c	ΔPA	program
			ms	arcsec		S	deg	
HIP 2472	HIP2472	2013-11-13		0.680	1.077	30x1x49.46	15.7	GS-ENG-GPI-COM
HIP 53524	HD95086	2013-12-10		0.330	1.356	21x1x119.29	15.0	GS-ENG-GPI-COM
HIP 64995	HD 115600	2014-04-22		0.365	1.157	58x1x49.46	32.5	GS-2014A-SV-403
HIP 11964	CC Eri	2014-11-08		0.865	1.057	36x1x59.65	31.9	GS-2014B-Q-500
HIP 12964	HR 826	2014-11-08		0.930	1.015	36x1x59.65	69.6	GS-2014B-Q-500
HIP 560	HR 9	2014-11-08		0.890	1.031	35x1x59.65	44.1	GS-2014B-Q-500
HIP 19893	gam Dor	2014-11-09		0.635	1.073	38x1x59.65	28.7	GS-2014B-Q-500
HIP 12413	HR 789	2014-11-09		0.785	1.039	40x1x59.65	43.5	GS-2014B-Q-500
HIP 490	HD 105	2014-11-09		0.615	1.026	40x1x59.65	31.2	GS-2014B-Q-500
HIP 25283	HD 35650	2014-11-09		0.650	1.023	40x1x59.65	24.9	GS-2014B-Q-500

Notes. a: coming from MASS measurements, not available before April 2015 and after April 2017. b: average between MASS and DIMM measurements; stale MASS measurements (non-zero values repeated over long - daily to monthly - periods of time; see Poyneer et al. 2016) were 1176 identified and removed. No MASS/DIMM values were available after April 2017. c: int. time = number of frames × number of co-added images \times Detector Integration Time per frame. Δ PA represents the parallactic rotation over the sequence.

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Appendix C: Companion candidates 1178

Table C.1: Companion candidates detected in this work. The table is available in electronic form at the CDS.

star name	obs. night	SNR	separation	PA	contrast	H2	H2-H3	algo.	classification
			arcsec	deg	mag	mag	mag		
HD 104467	2018-03-26	5.8	0.368 ± 0.005	7.5 ± 0.9	13.3	15.4 ± 0.1	0.3 ± 0.2	PACO	ambiguous
HD 118991A	2015-04-04	45.9	1.162 ± 0.002	217.7 ± 0.2	10.8	11.6 ± 0.1	0.3 ± 0.2	PACO	pm bkg
HD 131399A	2017-02-15	9.6	0.801 ± 0.003	193.9 ± 0.3	13.5	15.4 ± 0.1	0.4 ± 0.2	PACO	pm bkg
HD 131399A	2017-04-20	10.8	0.802 ± 0.003	194.0 ± 0.3	13.2	15.1 ± 0.1	0.4 ± 0.3	PACO	pm bkg
HD 24072	2018-11-22	—	0.193 ± 0.002	16.5 ± 0.4	6.9		—	cADI	star comp
HD 24072	2018-11-22	5.8	0.466 ± 0.004	337.1 ± 0.5	14.1	14.3 ± 0.7	$-2.0^{+0.1}_{-0.3}$	PACO	interesting
HD 36869	2016-12-17	17.2	0.783 ± 0.002	212.6 ± 0.2	12.7	15.9 ± 0.1	0.1 ± 0.1	PACO	pm bkg
HD 36869	2016-12-17	15.0	0.432 ± 0.002	107.5 ± 0.3	12.1	15.3 ± 0.1	0.0 ± 0.1	PACO	pm bkg
HD 74341B	2015-12-20	10.4	1.454 ± 0.007	342.3 ± 0.4	13.8	16.3 ± 0.1	0.0 ± 1.2	PACO	cmd bkg
HD 84330B	2015-12-18	20.7	1.123 ± 0.003	247.5 ± 0.3	12.8	15.7 ± 0.1	0.0 ± 0.1	PACO	pm bkg
HD 84330B	2016-03-18	11.8	1.072 ± 0.003	246.0 ± 0.3	13.1	16.0 ± 0.1	0.1 ± 0.4	PACO	pm bkg

Notes. H2: absolute SPHERE H2 magnitude. H2 - H3: SPHERE H2-H3 color. Classification: cmd bkg = background star via CMD; pm bkg: 1179 background star via proper motion analysis; sub comp: substellar companion; star comp: stellar companion. ^a: unconfirmed, see Section D.

1181 Appendix D: Stellar companions

In addition to the brown dwarfs HD 984 B and PZ Tel B, the 1182 ADI reduction identified 7 bright companion candidates (two 1183 of them detected twice in two different epochs). Proper motion 1184 analysis allowed us to identify one of them (namely, the one seen 1185 next to HIP 61087) as a background object and one as a bound 1186 companion (around HIP 74696). For the remaining objects, for 1187 which only one observation was available in our sample, we 1188 searched for archival detections in the literature. It turns out that 1189 1190 all the candidates but one (around HD 74341B) had been already 1191 imaged in the course of past campaigns, but just one (around HIP 38160) had already been confirmed as a comoving object 1192 1193 through follow-up observations (Rameau et al. 2013). Therefore, we performed the proper motion analysis for all the systems, us-1194 ing the astrometric measurements reported in Table D.1. 1195

The proper motion test confirmed that the 5 sources with two epochs exhibit a significantly different motion compared to static background objects, with large displacements related to orbital motion (Figure D.1).

In order to clarify the status of HD 74341 B, to further cor-1200 roborate the bound nature of the other objects and, finally, to 1201 test the reliability of the derived photometric masses, we ran 1202 the GaiaPMEX tool (Kiefer et al. 2024) to see if the astrome-1203 try of the primary from Gaia and/or Hipparcos showed hints 1204 1205 of wobbles indicative of the presence of an unseen companion. GaiaPMEX comes equipped with a model of the Renormal-1206 ized Unit Weight Error (ruwe; see Lindegren et al. 2021) and of 1207 the Gaia-Hipparcos proper motion anomaly (PMa; see Kervella 1208 et al. 2019, 2022) distribution expected for a single star as a func-1209 tion of stellar magnitude and colors. The evaluation of whether 1210 astrometric information is consistent with an unseen compan-1211 ion is performed in the following way. After defining a log-1212 uniform grid of companion masses $M_c \in [0.1, 3000] M_{Jup}$ and 1213 semi-major axes $a \in [0.01, 1000]$ au with 30×30 bins, the pro-1214 gram draws, within each bin, 100 (log M_c , log a)-doublets from 1215 a uniform distribution. As regards the other orbital parameters, 1216 they are randomly extracted from the distributions described in 1217 Table D.2. We employ stellar parallaxes from Gaia DR3, while 1218 stellar masses are recovered from our analysis described in Sec-1219 tion 2.3. 1220

At each node of the mass–sma grid, distributions of the ruwe and/or PMa are determined given the target and its hypothetical orbiting companion, and compared to the actual ruwe and/or PMa; the derivation of confidence regions for possible companion masses and semi-major axes can be finally obtained through Bayesian inversion.

1227 Clear astrometric detections were found for all the targets but 1228 HD 74341B (Figure D.2), due to the absence of the star in the 1229 Hipparcos catalog; due to the much shorter timespan of the as-1230 trometric measurements underlying the ruwe (\sim 3 yr, compared 1231 with the \sim 24 yr of the PMa), the sensitivity of the ruwe at the 1232 relatively large separation of the companion candidate is virtu-1233 ally null.

We find good agreement between the photometric and the dynamical masses for three stellar companions (those around HIP 74696, HD 24072, and HIP 26369). As regards the companions to HIP 67199 and HIP 38160, which are the ones located at the shortest separations from the star, we confirm their bound nature but we find largely underestimated masses, despite the expedient to use no-ADI instead of cADI¹⁷. We attribute the discrepancy to the fact that these sources lie at the edge of the 1241 coronagraph, where the transmission is much lower than elsewhere across the field of view. 1243

¹⁷ The no-ADI algorithm can be thought as a cADI but with no median-subtraction step. The advantage of the method is to avoid self-subtraction of signal from the source, a problem becoming more severe

at shorter separations; on the other hand, this is obtained at the price of much poorer detection limits.

Table D.1: Stellar companions identified in the sample with their astrometric and photometric properties. In addition to GPI measurements, literature astrometry is reported too.

STAR	DATE	SEP	PA	CONTRAST	ruwe	PMa	Zruwe	Z _{PMa}	MASS	source
		mas	deg	mag		mas yr^{-1}			$M_{ m Jup}$	
LID 67100	2015-04-04	114 ± 6	354.3 ± 2.9	$8.3^{a,b}$	1.02	6.10 ± 0.03	0.2	37 /	$>20^{\circ}$	TW
1111 0/199	2019-03-07	147.28 ± 0.17	51.41 ± 0.06	_	1.02	0.10 ± 0.03	0.2	57.4	>20	W23b
HD 74341B	2015-12-20	744 ± 2	75.9 ± 0.3	4.3	0.74	—	1.5	—	530 ± 45	TW
HIP 26360	2018-01-06	155 ± 1	222.5 ± 0.6	3.9	3.80	10.40 ± 0.22	28.1	21.8	125 ± 35	TW
1111 20509	2017-01-16	284.71 ± 0.3	201.78 ± 0.06	_	5.00	10.40 ± 0.22	20.1	21.0	125 ± 55	B22
HD 24072	2018-11-22	193 ± 2	16.5 ± 0.4	5.3	1 07		12.2		355+95	TW
11D 24072	2017-12-02	124.97 ± 0.77	7.81 ± 0.35	_	1.97	_	12.2		333 ₋₁₀₀	B22
LID 38160	2015-04-08	128 ± 6	283.4 ± 2.5	5.6^{b}	1 3/	30.65 ± 0.00	27	023	$>240^{\circ}$	TW
1111 38100	2009-11-25	141 ± 13	117.08 ± 2.28	—	1.54	50.05 ± 0.09	2.7	92.5	>240	R13
	2015-07-29	156 ± 2	357.8 ± 0.6	5.0						TW
HIP 74696	2019-08-11	139 ± 6	25.3 ± 2.3	5.3	0.87	5.10 ± 0.03	1.3	28.9	380^{+52}_{-50}	TW
	2023-04-19	118.94 ± 0.18	56.94 ± 0.07						50	W23a

Notes. ^{*a*}: no-ADI contrast; ^{*b*}: upper limit; ^{*c*}: lower limit. W23a: Waisberg et al. (2023a); W23b: Waisberg et al. (2023b); B22: Bonavita et al. (2022); R13: Rameau et al. (2013); TW: this work.



Fig. D.1: Proper motion test for the five stellar companion candidates with multiple epochs. As in Figure 4, a filled star indicates the displacement expected for a bound source with no relative motion to the star, whereas an empty star marks the location of a static background source. Second epochs are labeled by a '2', third epochs by a '3'. The large deviations from the filled star are likely caused by orbital motion.



Fig. D.2: Upper panel: flux maps (in contrast units) showing the stellar companion candidates detected with cADI. Lower panel: GaiaPMEX (sma, mass) maps, with contours outlining the area corresponding to the 68% and 95% confidence level. Photometric masses (dots) or lower limits (arrows) are overplotted for comparison. The HD 74341B map should be interpreted as a nondetection, the white area being incompatible with the absence of a signal.

Table D.2: Physical and or	bital parameters used in GaiaPMEX.
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Parameter	type	bounds or law
$\log M_c$	uniform	$\log M_c \pm \Delta \log M_c$
$\log a$	uniform	$\log a \pm \Delta \log a$
е	uniform	0-0.9
ω	uniform	$0-\pi$
Ω	uniform	$0-2\pi$
ϕ	uniform	0-1
I_c	$\sin I_c$ uniform	$0 - \pi/2$
$\overline{\omega}$	normal	$\mathcal{N}(arpi, \sigma_{arpi}^2)$
M_{\star}	normal	$\mathcal{N}(M_{\star}, \sigma^{\tilde{2}}_{M_{\star}})$

Notes. *e*: eccentricity; ω : periastron longitude; Ω : longitude of ascending node; ϕ : phase; I_c : inclination; ϖ : parallax; M_* : stellar mass.

1244 Appendix E: Effect of input assumptions on

1245 occurrence rates

We explored the dependence of the results derived in Section 3 on several input assumptions: the uncertainty on stellar age, the choice of the substellar evolutionary model, the degree of disequilibrium chemistry of planet atmospheres. In principle, all of them are expected to induce systematic deviations in the luminosity-mass relation, possibly impacting the reliability of the derived occurrence rates.

1253 As a first step, we evaluated the impact of model selection by repeating the computations from Table 3 using the AMES-1254 Cond models (Baraffe et al. 2003) and the Sonora Bobcat mod-1255 els (Marley et al. 2021). AMES-Cond models ignore the effect 1256 of dust opacity and are therefore more appropriate for objects 1257 with $T_{\rm eff} \leq 1300$ K compared to fully dusty models such as the 1258 1259 AMES-Dusty models (Baraffe et al. 2003). The derived com-1260 pleteness maps are shown on the left side of Figure E.1; the dif-1261 ferences between completeness values are plotted on the right 1262 side. In this regard, we stress that, given that it is the mean de-1263 tection probability across the (mass, sma) area \mathcal{A} that enters into Eq. 2, absolute differences are a more accurate proxy than rel-1264 ative differences when evaluating the impact of completeness 1265 maps variations on the derived frequency posteriors. Inspection 1266 of Figure E.1 clearly indicates that the discrepancies are the 1267 widest in the mass range [1, 5] M_{Jup} , and rapidly decrease at 1268 larger masses: this can be seen as a consequence of the stronger 1269 cooling rate of less massive objects, combined with the larger 1270 theoretical uncertainties at lower masses. 1271

As a consequence of this observation, we expect the lower mass value selected to define \mathcal{A} , M_{low} to have a large impact on the accuracy of the results. We quantified this effect by computing occurrence rates under the three models for $M_{\text{low}} =$ [1, 2, 3, 4, 5] M_{Jup} . As expected, the problem exacerbates for lower values of M_{low} (Figure E.2). This test justifies our conservative choice for \mathcal{A} (Sec. 3.4).

Afterwards, we investigated the dependence of the results on 1279 the assumption of chemical equilibrium: in particular, we used 1280 two suites of ATMO models that assume 1) chemical equilibrium 1281 (ATMO-CEQ) or 2) strong chemical disequilibrium (ATMO-1282 NEQ-S), that is, a different relation for the vertical mixing coef-1283 ficient (Phillips et al. 2020). Given a certain H-band magnitude, 1284 the fractional mass difference, computed as a function of age 1285 and ATMO-NEQ-W mass ($m \in [1, 10] M_{Jup}$), can be as large as 1286 30% compared to the chemical equilibrium case. The variation 1287 is larger at lower masses and larger ages, that is, at lower effec-1288 tive temperatures. The derived completeness maps, analogous to 1289 Figure E.1, are shown in Figure E.3. Given our careful choice of 1290 mass boundaries, it is possible to argue that completeness values 1291 within \mathcal{R} are dominated by projection effects rather than detec-1292 tion limits. Hence, we expect the occurrence rates to be fully 1293 consistent with those of the original analysis. 1294

Finally, we provide similar completeness maps to quantify the dependence on age uncertainty: Figure E.4 shows the variation of the maps when assuming lower and upper values for stellar ages.

All the occurrence rates derived in this Section are visually compared in Figure E.5. It is evident that any doublet of estimates is compatible within the errors, making the estimates presented in this work robust against systematic effects.



Fig. E.1: Effect of model selection on survey completeness: maps assuming the Ames-COND model (top row) and the Sonora model (bottom row). Left panels show completeness maps, while right panels indicate the difference relative to the map used for the analysis. The green dashed box indicates our nominal choice of \mathcal{A} .



Fig. E.2: Trend of the uncertainties in the derived occurrence rates with M_{low} . Each model – shown as a triplet (nominal ages, lower ages, upper ages) – is plotted in a different color. Horizontal offsets have been applied to each line for the sake of visualization.

Article number, page 22 of 21



Fig. E.3: Effect of non-equilibrium chemistry on survey completeness: maps using the ATMO models assuming equilibrium chemistry (top row) and strong disequilibrium chemistry (bottom row). Left panels show completeness maps, while right panels indicate the difference relative to the map used for the analysis. The green dashed box indicates our nominal choice of \mathcal{A} .



Fig. E.4: Effect of age uncertainty on survey completeness: maps assuming lower (top row) and upper (bottom row) values for stellar ages. Left panels show completeness maps, while right panels indicate the difference relative to the map used for the analysis. The green dashed box indicates our nominal choice of \mathcal{A} .



Fig. E.5: Effect of model selection and age uncertainty on planet occurrence ($\mathcal{A} = [5, 13] M_{Jup} \times [10, 100]$ au): results for the entire sample (red squares), the BA subsample (blue diamonds), the FGK subsample (green circles) using: the standard ATMO-NEQ-weak model (ATMO-NW); the same model with lower (ATMO-NW, L) and upper (ATMO-NW, U) ages; the ATMO model with no (ATMO-C) and strong (ATMO-NS) disequilibrium chemistry; the AMES-Cond model (AMES-COND) and the Sonora Bobcat model (SONORA). A Jeffreys prior is assumed (see Section 3.4).