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1	Motor inhibition prevents motor execution during typing imagery: evidence from an
2	action-mode switching paradigm
3	Ladislas Nalborczyk ^{1,2,3} , FXavier Alario ¹ , & Marieke Longcamp ¹
4	¹ Aix Marseille Univ, CNRS, CRPN, Marseille, France
5	2 Cognitive Neuroimaging Unit, CEA DRF/I2BM, INSERM, Université Paris-Sud,
6	Université Paris-Saclay, NeuroSpin Center, Gif/Yvette, France
7	3 Inser m U 1127, CNRS UMR 7225, Sorbonne Universités, Institut du Cerveau, ICM,
8	Paris, France

Author Note

¹⁰ Correspondence concerning this article should be addressed to Ladislas Nalborczyk,

¹¹ CRPN, CNRS & Aix-Marseille University, 3 place Victor Hugo, 13331 Marseille Cedex 3,

12 France. E-mail: ladislas.nalborczyk@gmail.com

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Abstract

Motor imagery is accompanied by a subjective multisensory experience. This sensory 14 experience is thought to result from the deployment of internal models developed for the 15 execution and monitoring of overt actions. If so, how is it that motor imagery does not 16 lead to overt execution? It has been proposed that inhibitory mechanisms may prevent 17 execution during imagined actions such as imagined typing. To test this hypothesis, we 18 combined an experimental with a modelling approach. We conducted an experiment in 19 which participants (N = 49) were asked to alternate between overt (executed) and covert 20 (imagined) typing. We predicted that motor inhibition should lead to longer reaction and 21 movement times when the current trial is preceded by an imagined vs. an executed trial. 22 This prediction was borne out by movement times, but not by reaction times. We 23 introduced and fitted an algorithmic model of motor imagery to disentangle potentially 24 distinct inhibitory mechanisms underlying these effects. Results from this analysis suggest 25 that motor inhibition may affect different aspects of the latent activation function (e.g., the 26 shape of the activation function or the motor execution threshold) with distinct 27 consequences on reaction times and movement times. Overall, these results suggest that 28 typing imagery involves the inhibition of motor commands related to typing acts. 29 Preregistration, complete source code, and reproducible analyses are available at 30 https://osf.io/y9a3k/. 31

Keywords: typing, typing imagery, motor imagery, motor inhibition, response
 inhibition, algorithmic modelling

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Introduction

Motor imagery, defined as the mental rehearsal of an action, is a cornerstone of 36 human cognition. For most individuals, it is accompanied by a rich subjective multisensory 37 experience. A prominent proposal is that during motor imagery, the sensory consequences 38 of actions may be simulated mentally using pairs of internal models developed for the 39 control of overt actions (e.g., Grush, 2004; Jeannerod, 1994, 2001). More precisely, internal 40 forward models may predict the sensory consequences of (a copy of) motor commands 41 issued from internal inverse models. This view is supported by a wealth of chronometric, 42 electrophysiological, neuroimaging, and neurostimulation studies (for reviews, see Guillot, 43 Di Rienzo, et al., 2012; Guillot, Hoyek, et al., 2012; Guillot & Collet, 2005). 44

This simulationnist perspective entails that the reuse of motor networks during motor 45 imagery should be accompanied by mechanisms preventing execution (Jeannerod, 2006; 46 O'Shea & Moran, 2017). Previous research has postulated at least three (non-exclusive) 47 potential inhibitory mechanisms that might operate during motor imagery (Guillot, Di 48 Rienzo, et al., 2012). First, the need to prevent execution could be integrated within the 49 representation of the action to be produced internally so that only "subthreshold" motor 50 commands are involved during motor imagery (see also Bach et al., 2022; Glover et al., 51 2020). Second, motor inhibition could be applied broadly to all ongoing actions and to all 52 effectors involved in these actions (global motor inhibition). Third, motor inhibition may 53 be applied in a finer-grained manner, only to some effector (effector-specific inhibition). 54

To investigate these inhibitory mechanisms, Rieger et al. (2017) developed a protocol called the action-mode switching paradigm in which participants have to rapidly alternate between various overt (executed) and covert (imagined) hand movements. The authors reasoned that, if motor imagery involves the inhibition of motor commands, then the inhibition applied during some imagined trial may persist to the next trial and may slow down its initiation, resulting in switching costs (for imagined-executed

vs. executed-executed sequences) or switching benefits (for imagined-imagined 61 vs. executed-imagined sequences). By further varying the effector used for the executed or 62 imagined action, this paradigm allows assessing the presence of effector-specific inhibition. 63 Using this paradigm, Rieger et al. (2017) showed that imagined hand movements involve 64 both global and effector-specific motor inhibition (see also Scheil & Liefooghe, 2018). It has 65 been further shown that these effects do not depend on the effector used to indicate the 66 onset and duration of imagined movements (Bart et al., 2020). These effects are influenced 67 by the proportion of imagined trials in mixed blocks (Bart et al., 2021b) and they decay 68 rapidly with increasing inter-trial interval (Bart et al., 2021a). In brief, the existence of 69 inhibitory mechanisms preventing motor execution during imagined hand reaching 70 movements has been successfully demonstrated and replicated in several distinct 71 experiments using this paradigm. 72

However, the evidence so far comes from ad-hoc pointing movements; it is unclear 73 whether the observed effects and inferred mechanisms generalise to more intricate and 74 automatised actions. Among our daily activities, typing provides an ideal test case of this 75 generalisation. Typing involves complex and often highly automatised sequential motor 76 actions, and it is ubiquitous in the everyday life of millions of persons. Cognitive models of 77 typing propose that all keystrokes programs of a word are activated in parallel prior to the 78 onset of execution, with a graded activation level allowing the ordering of the keystrokes 79 during execution (Logan & Crump, 2011; Pinet et al., 2016; Pinet et al., 2019; Rumelhart 80 & Norman, 1982; Snyder et al., 2014). Moreover, it has been hypothesised that during 81 typing imagery, forward models may predict the sensory consequences of (a copy of) typing 82 acts issued from inverse models (Dahm & Rieger, 2019). In support of the view of typing 83 imagery as mentally simulated typing, previous research suggests that the timing of typing 84 imagery is shorter but proportional to the timing of overt typing (Rieger, 2012) and that 85 typing imagery generally contains similar errors as overt typing, albeit to a lesser extent 86 (Dahm & Rieger, 2019; Rieger et al., 2011). 87

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Here, we assessed the presence and scope (i.e., global and/or effector-specific) of 88 motor inhibition during typing imagery with an adapted version of Rieger et al. (2017)'s 89 action-mode switching paradigm. We asked participants to alternate between overt and 90 covert typing of unimanual words. We reasoned that, if typing imagery involves the 91 inhibition of motor commands (and if this inhibition persists until the next trial), then, 92 compared to typing execution, typing imagery should slow down the initiation of the next 93 trial (may this trial be imagined or executed). This should translate into a *switching cost* 94 in imagined-executed vs. executed-executed sequences, and into a *switching benefit* in 95 imagined-imagined vs. executed-imagined sequences. In contrast, if typing imagery does 96 not involve the inhibition of motor commands, we expected to observe a standard 97 switching cost (Kiesel et al., 2010; for reviews, see Monsell, 2003) in both executed and 98 imagined trials. Following Rieger et al. (2017), we hypothesised that effector-specific 99 inhibition should translate into longer reaction times (RTs) and movement times (MTs) 100 when the same hand is repeated than when it is not repeated, only in sequences in which 101 the first trial is an imagery trial. 102

Differences in RTs or MTs due to the action mode of the previous trial provide 103 important cues regarding the inhibitory mechanisms involved. However, these raw 104 contrasts may be inconclusive about the mechanistic implementation. For instance, motor 105 imagery could slow down the initiation of subsequent movements either by upregulating 106 the motor execution threshold or by delaying the spread of excitatory inputs (Nalborczyk 107 et al., 2024). In order to disentangle potentially distinct underlying inhibitory mechanisms, 108 we fitted a novel algorithmic model of motor imagery allowing to infer the effect of 109 inhibition in the previous trial on the underlying timecourse of motor activation during 110 typing imagery in the current trial (see the modelling section). 111

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Methods

In this section, we report how we determined our sample size, all data exclusions (if any), all manipulations, and all measures in the study (Simmons et al., 2012). A pre-registered version of our protocol is available online: https://osf.io/y9a3k/.

116 Ethics information

The present research has been approved by the local ethics committee of Aix-Marseille University (agreement "HowFast", number 2016-09-11-06). All participants provided informed consent and received course credits in exchange for their participation.

120 Participants

¹²¹ We recruited 49 French-speaking undergraduate students in Psychology from ¹²² Aix-Marseille University, ranging in age from 18 to 23 years (M = 19.41, SD = 1.32, 39F, ¹²³ 10M), and with no reported history of psychiatric or neurological disorder, speech disorder, ¹²⁴ or hearing deficit. As preregistered (https://osf.io/x5r67), this sample size was defined ¹²⁵ based on temporal constraints (i.e., two full weeks of data collection) and previous research ¹²⁶ (i.e., more than double the number of participants in Rieger et al., 2017).

127 Linguistic material

Two classes of words were created based on the location of their constitutive letters 128 on the keyboard (in relation to a median line located between the t-g-b and y-h-n letters 129 on a regular AZERTY/QWERTY keyboard). Words made of letters located on the left of 130 this median line were considered as "left-hand words" whereas words made of letters 131 located on the right of this median line were considered as "right-hand words". Left-hand 132 words included words such as "averse" (rainfall) or "carafe" (carafe), whereas right-hand 133 words included words such as "poumon" (lung) or "nylon" (nylon). These words were 134 taken from the Lexique database (New et al., 2004). The groups of left-hand and 135

right-hand words were matched for per-word average bigram frequency, word frequency,
number of letters, and number of syllables (the complete list of stimuli is available in the
online supplementary materials).

139 Design

The experimental design was fully within-participant, with three crossed two-level factors: current action mode (i.e., executed vs. imagined trials), previous action mode (i.e., executed vs. imagined trials), and hand alternation/repetition (i.e., same hand vs. other hand), defining a total of eight sequences (i.e., pairs of successive trials) of interest.

144 Procedure

At the beginning of the experiment, typing expertise was assessed via an online copy task (Leijten & Van Waes, 2013; Van Waes et al., 2019) and summarised by computing average typing speed (ranging from 171 to 470.80 characters per minute, M = 313.25, SD = 65.13) and accuracy (ranging from 88% to 99% of correctly typed characters, M = 94.52, SD = 2.77). We also assessed typing habits via the questionnaire developed in Pinet et al. (2022) (see the online supplementary materials).

Afterwards, we provided participants with extensive instructions about typing imagery. Namely, we instructed them to imagine themselves from a first-person perspective typing the words, insisting on the multisensory nature of motor imagery. We asked them to focus on the tactile feelings and auditory percepts associated with typing (i.e., kinaesthetic motor imagery). Participants were also instructed to perform both the overt and covert typing tasks as fast and accurately as possible without correcting potential errors.

Participants started with a first training block which consisted only of overt typing trials. They continued with a second training block consisting only of covert typing trials followed by a third training block containing both overt and covert typing trials. Within each trial, the action mode (i.e., executed vs. imagined trials) was indicated by a geometrical shape (i.e., rectangle vs. circle) surrounding the word. The association between
action mode and shape was counterbalanced across participants. The main experimental
part (post-training) followed, consisting of six blocks of 65 trials each, yielding a total of
390 trials per participant (Figure 1A). The order of word sequences (i.e., pairs of successive
words/trials) in each block was randomised using Euler tours (Bakermans & Behrens,
2021), ensuring that each of the eight trials sequences appeared equally often (i.e., 8 times
per block).



Figure 1. A: Experimental procedure. The main experimental part (post-training) involved6 blocks of 65 trials each. B: Illustrated timecourse of a single trial.

The timecourse of each single trial is depicted in Figure 1B. Participants first needed to press the spacebar with their two thumbs for at least 500ms, visually represented by a progress bar displayed at the center of the screen. After these 500ms, the word appeared on the screen until the participant released the spacebar and started typing the word, either overtly or covertly. As soon as they finished, they had to return their thumbs to the spacebar. In each trial, we measured both the reaction time (RT) and the movement time (MT). The reaction time was defined as the time interval between the appearance of the stimulus (i.e., the word) and the release of the spacebar by the participant (i.e., the disappearance of the stimulus from the screen). The movement time was defined as the time interval between the release of the spacebar by the participant and the next spacebar keypress (Figure 1B).

The first trial of each block was discarded (as it was not preceded by any other trial), yielding a total of 48 repetitions of each of the eight sequences of interest per participant. Following these six blocks, participants had to fill out the typing habits questionnaire developed in Pinet et al. (2022). The experimental procedure was developed using the PsychoPy software (Peirce et al., 2019) and took approximately 45 min. At the end of the experiment, participants were fully informed about the theoretical rationale for the study and compensated (in course credits) for their participation.

186 Data analysis

Statistical analyses were conducted using R version 4.2.3 (2023-03-15) (R Core Team, 187 2017) and are reported with the papaja (Aust & Barth, 2017) and knitr (Xie, 2015) 188 packages. To assess the effects of motor inhibition on RTs and MTs, we built and fitted 189 several Bayesian multilevel generalised linear models using the **brms** package (Bürkner, 190 2017).¹ Data were analysed using *current action mode* (2 levels, executed vs. imagined, 191 recoded using a -0.5/+0.5 sum contrast), previous action mode (2 levels, executed 192 vs. imagined, recoded using a -0.5/+0.5 sum contrast), and hand alternation/repetition (2) 193 levels, same vs. different, recoded using a -0.5/+0.5 sum contrast) as within-subject 194 categorical predictors, and the RT or MT as a dependent variable. We analysed RTs and 195 MTs separately (i.e., we built separate models for each of these two measures). All models 196 allowed intercepts and slopes to vary by participant. 197

¹ An introduction to Bayesian statistics is outside the scope of this paper. The interested reader is referred to Nalborczyk et al. (2019) for an introduction to Bayesian multilevel modelling using the brms package.

Models were fitted using weakly informative priors (see the supplementary materials 198 for code details). Four Markov Chain Monte-Carlo (MCMC) were ran for each model to 199 approximate the posterior distribution, including each 5000 iterations and a warmup of 200 2000 iterations. Posterior convergence was assessed examining trace plots as well as the 201 Gelman-Rubin statistic \hat{R} . Constant effect estimates were summarised via their posterior 202 mean and 95% credible interval (CrI), where a credible interval can be considered as the 203 Bayesian analogue of a classical confidence interval. When applicable, we also report Bayes 204 factors (BFs), computed using the Savage-Dickey method, which consists in taking the 205 ratio of the posterior density at the point of interest divided by the prior density at that 206 point. These BFs can be interpreted as an updating factor, from prior knowledge (what we 207 knew before seeing the data) to posterior knowledge (what we know after seeing the data). 208

As pre-registered, we excluded trials in which participants performed the wrong action mode, that is, trials in which participants typed some letters in the imagined typing condition or did not type in the executed typing condition. This amounted to 149 trials, that is, less than 0.01% of the total number of trials.

213

Results

This section is divided into four parts. First, we present a visual exploration of the 214 data. Second, we present results from confirmatory (preregistered) analyses, aiming at 215 assessing the sequential effects of motor inhibition on reaction times (RTs) and movement 216 times (MTs). Third, we present results from exploratory (non-preregistered) analyses, 217 aiming at distinguishing participants that effectively performed the imagery task from 218 those who did not. These were motivated by various indications that the task may not 219 have been performed equally well by all participants. Fourth, we discuss results obtained 220 by fitting a novel algorithmic model of motor imagery to these data and how it can be used 221 to disentangle different underlying inhibitory mechanisms. 222

²²³ Visual exploration of the data

Figure 2 shows the distributions of RTs, MTs, and total response times (TTs) for the different sequences of trials (i.e., pair of successive trials): imagined-imagined, executed-imagined, imagined-executed, and executed-executed sequences (from left to right). Given this visual arrangement, if imagined typing involves motor inhibition, this should translate into negative slopes from left to right columns, in each panel. In short, such general pattern was observed for MTs (and TTs), whereas the inverse pattern was observed for RTs.

²³¹ Confirmatory (preregistered) analyses

To estimate these effects while accounting for the skewness of the collected data (for 232 more details, see the online supplementary materials), we fitted two multilevel 233 distributional Log-Normal models (one model for RTs and one model for MTs), where 234 "distributional" means that not only the means but also the standard deviations were 235 allowed to vary across conditions (for more details, see for instance Williams et al., 2021). 236 This was justified by the observation that the amount of variability (both within and 237 between participants) strongly differed across executed and imagined trials (as can be seen 238 from Figure 2). The distributional model allows estimating the effects of interest while 239 taking into account the effects the predictor variables may have on dispersion, thus 240 producing more precise (less biased) estimates of the effects on means or medians. 241 Estimates from these models are reported in Table 1 and Table 2 and discussed in the next 242 two paragraphs. 243

Reaction times. RTs were slightly shorter when the previous trial was imagined rather than executed ($\beta = -0.006, 95\%$ CrI [-0.013, 0], BF₁₀ = 1.65, BF₊ = 0.036) and longer when the current trial was imagined rather than executed ($\beta = 0.066, 95\%$ CrI

Table 1

Predictor	Estimate	SE	Lower	Upper	Rhat	BF_{10}	BF_+
previous mode	-0.006	0.004	-0.013	0.000	1.001	1.65	0.036
current mode	0.066	0.004	0.059	0.073	1.000	6×10^{18}	∞
same hand	-0.004	0.003	-0.011	0.003	1.001	0.666	0.151
previous mode:current mode	-0.015	0.006	-0.027	-0.003	1.000	11.985	0.006
previous mode:same hand	-0.001	0.006	-0.013	0.010	1.000	0.614	0.665
current mode:same hand	-0.007	0.006	-0.018	0.005	1.000	1.107	0.154
previous mode:current mode:same hand	0.004	0.008	-0.012	0.021	1.000	0.946	2.360

Estimates from the multilevel Log-Normal model fitted on reaction times

Note. The 'Estimate' column represents the estimated group-level effect (slope) of each predictor included in the model (in terms of standardised AUCs). The 'Lower' and 'Upper' columns contain the lower and upper bounds of the 95% CrI, whereas the 'Rhat' column reports the Gelman-Rubin statistic. The last two columns report the BF in favour of the alternative hypothesis (relative to the null) and the directional (i.e., one-sided) BF, respectively.

 $[0.059, 0.073], BF_{10} = 6 \times 10^{18}, BF_{+} = Inf).^{2}$ There was strong evidence in favour of a 247 non-null interaction effect between the effect of the previous mode and the effect of the 248 current mode ($\beta = -0.015, 95\%$ CrI [-0.027, -0.003], BF₁₀ = 11.985, BF₊ = 0.006), 249 indicating that the effect of the previous mode was stronger in imagined trials than in 250 executed trials (where the effect of the previous trial was null). This equates to a switch 251 cost in executed-imagined compared to imagined-imagined sequences, whereas there was no 252 evidence for such an effect in imagined-executed compared to executed-executed sequences. 253 These results are more readily understandable visually and are presented in the original 254 scale of the RTs in Figure 3. 255

 $^{^{2}}$ For one-sided hypotheses, BF₊ represents the ratio of the posterior probability of the effect being positive and the posterior probability of the effect being negative.

Table 2

Predictor	Estimate	SE	Lower	Upper	Rhat	BF_{10}	BF_+
previous mode	0.016	0.004	0.009	0.023	1.000	1041.083	23999.000
current mode	0.037	0.004	0.029	0.045	1.000	5×10^{15}	∞
same hand	0.001	0.004	-0.006	0.008	1.000	0.382	1.420
previous mode:current mode	-0.011	0.006	-0.023	0.001	1.000	2.97	0.036
previous mode:same hand	0.000	0.006	-0.012	0.012	1.000	0.61	0.904
current mode:same hand	-0.005	0.006	-0.017	0.007	1.000	0.826	0.282
previous mode:current mode:same hand	0.007	0.008	-0.009	0.024	1.000	1.195	4.079

Estimates from the multilevel Log-Normal model fitted on movement times

Note. The 'Estimate' column represents the estimated group-level effect (slope) of each predictor included in the model (in terms of standardised AUCs). The 'Lower' and 'Upper' columns contain the lower and upper bounds of the 95% CrI, whereas the 'Rhat' column reports the Gelman-Rubin statistic. The last two columns report the BF in favour of the alternative hypothesis (relative to the null) and the directional (i.e., one-sided) BF, respectively.

Movement times. MTs were longer when the previous trial was imagined rather 256 than executed ($\beta = 0.016, 95\%$ CrI [0.009, 0.023], BF₁₀ = 1041.083, BF₊ = 23999) and 257 longer when the current trial was imagined rather than executed ($\beta = 0.037, 95\%$ CrI 258 $[0.029,\,0.045],\,\mathrm{BF}_{10}=5\times10^{15},\,\mathrm{BF}_+=\mathrm{Inf}).$ This equates to a switching benefit in 259 executed-imagined compared to imagined-imagined sequences, and to a switching cost in 260 imagined-executed compared to executed-executed sequences. There was moderate 261 evidence in favour of a non-null interaction effect between the effect of the previous mode 262 and the effect of the current mode (β = -0.011, 95% CrI [-0.023, 0.001], BF_{10} = 2.97, BF_+ 263 = 0.036), suggesting that the effect of the previous mode was smaller in imagined trials 264 than in executed trials. These results are also depicted in Figure 3. 265

The predictions regarding effector-specific inhibition could not be appropriately tested, as highlighted by a reviewer. The participants included were not necessarily touch typists. Therefore, some letters may have been typed with either the left or the right hand according to the context or to individual preferences rather than to their keyboard position; words containing central letters on the keyboard (e.g., T, G, or B) may have been typed with both hands. This uncertainty blurs the experimental manipulation we implemented and prevents us from drawing conclusions regarding effector-specific inhibition effects. For the sake of completeness, though, we report the results from these analyses in the supplementary materials.

Preliminary summary. To sum up, results from these analyses revealed that, as
predicted, imagery in the previous trial increased the *duration* of executed or imagined
movements, suggesting that typing imagery does involve the inhibition of motor
commands. However, contrary to our predictions and results from previous studies
studying pointing movements (e.g., Bart et al., 2020, 2021a, 2021b; Rieger et al., 2017), the *initiation* of executed or imagined movements (i.e., the RT) was sped up rather than
slowed by typing imagery in the previous trial.

²⁸² Exploratory (non-preregistered) analyses - Assessing task compliance

In this section, we report the results of exploratory (i.e., non-preregistered) analyses, 283 aiming at distinguishing participants who effectively performed the imagery task from 284 those who did not. How can we make sure that participants effectively performed the task 285 as expected in imagined trials? This is a common concern in studies involving covert 286 operations such as mental or motor imagery. Although there is no possible direct 287 verification, we considered an ensemble of a priori and a posteriori precautions. A priori, 288 we took great care of explaining the task to the participants in details and planned an 289 extensive training period during which the experimenter could monitor the (overt) 290 behaviour of the participant to provide more details about the task, if needed. 291

A posteriori (after data collection), we sought to separate participants who were plausibly engaged in the imagery task from those who were not, based on the collected

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movement times. We reasoned that if the imagined action follows a "faithful" mental simulation of the corresponding overt action, we could expect the time it takes to imagine an action to be correlated to the time it takes to execute the same action (similar to what has been observed for handwriting or walking, for review see Guillot & Collet, 2005).

Figure 4 depicts the relation between the average movement time in executed and 298 imagined trials, averaged per word and participant, for a few exemplary participants. In 299 the data, we can distinguish between two profiles of participants: those who show a 300 positive relation between executed and imagined movements times (left panel, in green) 301 and those who do not (right panel, in orange), thereafter labelled as "on-task" and 302 "off-task" participants, respectively. We can then construct groups of participants (of 303 similar size) according to the correlation between the executed and imagined movement 304 times and assess the effect of interest according to this new grouping factor. 305

Figure 5 shows that the predicted effects of the previous action mode on MTs was 306 only found in participants with a non-null correlation between executed and imagined MTs 307 (middle and right panels) and was the strongest in participants with a strong positive 308 correlation between executed and imagined MTs (right panel). This visual intuition was 309 confirmed by fitting the model described in the previous section and adding interaction 310 terms with the correlation coefficient. This analysis revealed a strong interaction between 311 the effect of the previous action mode and the correlation coefficient, suggesting that higher 312 correlation coefficients (between executed and imagined MTs) were associated with higher 313 slowing effects of the previous action mode. This strongly suggests that typing imagery, 314 when performed correctly, does involve the inhibition of motor commands. However, this 315 figure also shows that even for "on-task" participants (right panel), we did not find the 316 effect predicted based on previous studies for the RTs. Finally, this analysis revealed that 317 participants with the strongest positive correlation coefficients had longer RTs on average. 318

³¹⁹ Modelling latent inhibitory processes

Differences in RTs or MTs due to the action mode of the previous trial provide 320 important cues regarding the inhibitory mechanisms involved. However, these raw 321 contrasts may be inconclusive about the mechanistic implementation. For instance, motor 322 imagery could slow down the initiation of subsequent movements either by upregulating 323 the motor execution threshold or by delaying the spread of excitatory inputs (Nalborczyk 324 et al., 2024). We developed an algorithmic model intended to capture alternative inhibitory 325 mechanisms, and to distinguish between modulations of the motor execution threshold and 326 modulations in the shape of the underlying activation function (Nalborczyk et al., 2024). 327 This model provides a simplified overarching description of how the motor system is 328 involved over the timecourse of a trial during motor imagery, roughly corresponding to the 329 underlying activity of populations of excitatory and inhibitory neurons (adapted from 330 MacDonald et al., 2014, 2017).³ The overall level of motor activation is modelled as a 331 time-varying rescaled lognormal function defined as: 332

$$f(t;A,\mu,\sigma) = A \cdot \exp\left[-\frac{(\ln t - \mu)^2}{2\sigma^2}\right], \quad t > 0$$

where A represents the amplitude (i.e., the maximum value) of the activation function and is usually fixed to an arbitrary value (e.g., A = 1). This leaves three free parameters: μ (the peak latency or "peak time" of the function), σ (its width or "curvature"), and one

³ In the initial stages of the model development, we thought that this model could be used to account for both motor imagery and execution. However, it soon became clear that this model could not be applied as it is for motor execution because in such cases it leads to nonsensical predictions. For instance, assuming both a motor imagery threshold and a motor execution threshold make sense for motor imagery. However making the same assumption for motor execution would imply a stage of conscious (experienced) motor imagery preceding execution, which we think is not cognitively plausible. Primarily for this reason, we introduced this model as a model of motor imagery alone and therefore fitted it to imagined trials only.

parameter for the imagery threshold T (expressed relative to A and such than T < A). From a psychological perspective, A - T may be related to the vividness of motor imagery percepts, μ may be related to the speed at which these mental percepts are established, and σ may be related to their duration (Nalborczyk et al., 2024). In this model, the reaction time is defined as the time at which the activation function crosses the threshold for motor imagery and the movement time is defined as the time "spent" above the threshold (Figure 6) (similar to what has been proposed for conscious access, e.g., Pereira et al., 2022).

We fitted this model to data coming from each participant individually. The fit procedure was adapted from Ratcliff and Smith (2004) and is commonly used to fit sequential sampling models to behavioural data (e.g., Servant et al., 2019). The model was simultaneously fitted to RTs and MTs quantiles (0.1, 0.3, 0.5, 0.7, 0.9), to take into account the entire RTs and MTs distributions (instead of only the mean or median values). The following loss function (likelihood-ratio chi-square statistic) was minimised:

$$G^2 = 2\left(\sum_{i=1}^2\sum_{j=1}^6 p_{ij}\log\frac{p_{ij}}{\pi_{ij}}\right)$$

where the outer summation over i refers to RTs and MTs, and the inner summation 349 over j extends refers to the 6 bins bounded by RT/MT quantiles. The quantities p_{ij} and 350 π_{ij} are the observed and predicted proportions of trials in RT/MT bin j. This G^2 statistic 351 characterises the goodness-of-fit of the model to the joint distributions of RT and MT. It 352 was minimised using differential evolution (Storn & Price, 1997) and 200 simulated trials 353 per condition (with F = 0.9 and CR = 0.95) using the DEoptim package (Mullen et al., 354 2011). The code underlying this model and the fitting routine are implemented in the 355 momimi package (available on Github: https://github.com/lnalborczyk/momimi). 356

Table 3

Average (group-level) parameter estimates from the algorithmic model of motor imagery

Condition (trial sequence)	Execution threshold	Peak time	Curvature
executed_imagined	1.34	1.23	0.68
imagined_imagined	1.24	1.22	0.68

Estimates from these models are reported in Figure 6 and Table 3. Recall that 357 whereas in previous studies both RTs and MTs increased following imagined trials, in the 358 present study group-level results showed that MTs increased but RTs decreased following 359 imagined trials. Parameter estimates obtained by fitting the algorithmic model to these 360 data suggest that these differences may be due to different underlying inhibitory 361 mechanisms. More precisely, the shortening of RTs (jointly observed with the lengthening 362 of MTs) following imagined trials is compatible with a *decrease* in the motor execution 363 threshold (as reported in Table 3), whereas the lengthening of both RTs and MTs following 364 imagined trials is compatible with a delay in the peak time of the activation function (as 365 previously suggested by Nalborczyk et al., 2024). 366

367

Discussion

We investigated the presence of motor inhibition during typing imagery with an 368 adapted version of Rieger et al. (2017)'s action-mode switching paradigm. Overall, our 369 results show that typing imagery slows down the realisation of subsequent executed or 370 imagined movements (as assessed by differences in MTs), but speeds up the initiation of 371 subsequent executed or imagined movements (as assessed by differences in RTs). Taken 372 together, these results suggest that motor inhibition prevents the execution of typing acts 373 during typing imagery. The algorithmic modelling revealed a possible implementation of 374 inhibition in terms of threshold modulation or delay. In the following, we discuss this 375

interpretation in more detail, including potential explanations for the discrepancies between
our results and previous results obtained using the action-mode switching paradigm.

³⁷⁸ Unexpected effects on reaction times

Based on previous studies using a similar paradigm with imagined and executed 379 pointing movements, we hypothesised that motor inhibition during imagined movements 380 may slow down both the RT and MT. However, we did not observe such an effect on the 381 RTs. The discrepancy between our results and those of previous studies may come from the 382 nature of the task(s) participants had to realise during the time period considered as the 383 "reaction time" in our experiment. During this time period, participants had to identify 384 the action of the current trial (as indicated by the geometrical shape surrounding the word 385 displayed on the screen), but they also had to read the word, and possibly process it at 386 different linguistic stages, which differs considerably from the RT period of previous studies 387 involving simpler hand movements and during which participants only had to identify the 388 spatially congruent target to which they should next reach (or imagined reaching). We 380 assessed this possibility in a second follow-up preregistered study (cf. supplementary 390 materials). The overall procedure was similar to the main experiment, except for the 391 linguistic material. In this experiment, we sought design an experiment closer to the 392 protocol of Rieger et al. (2017). To this end, we used bigrams instead of words while 393 taking care of maximising the distance between left-hand and right-hand bigrams on the 394 keyboard. Moreover, we added an association training during which participants learned to 395 associate four bigrams to a geometrical shape. This was intended to remove or reduce the 396 effect of reading the item on screen during the period considered as "reaction time" and, as 397 such, to provide a more direct comparison to the procedure of Rieger et al. (2017). Overall, 398 we observed no difference across sequences in average RTs or MTs, but we observed longer 399 RTs and MTs variability following imagined trials, in a direction compatible with the 400 expected effects on average RTs and MTs (cf. supplementary materials, Table S1 and S2). 401

Another possibility, suggested by a reviewer, is that cue repetition (i.e., the geometric 402 shape indicating the action mode) in successive trials of the same action mode may have 403 provided a perceptual benefit, which could counteract the hypothesised effect (especially on 404 RTs). In other words, the effect of task-switching may be confounded with the effect of 405 cue-switching (e.g., Schneider & Logan, 2011). Whereas this could explain why we did not 406 observe the predicted effect on RTs in II vs. EI sequences (where the effect of cue repetition 407 and action mode switching may cancel out), this could not explain the lack of effects on 408 RTs in IE vs. EE sequences. Indeed, in executed trials, perceptual benefits related to cue 400 (action mode) should be congruent (in the same direction) with inhibition effects and 410 should therefore results in longer RTs for IE vs. EE sequences, which is not what was 411 observed. Nevertheless, the effect of cue-switching could be assessed in a future extension 412 of the present study by using a 2:1 cue-task mapping to assess the impact of this 413 manipulation on the pattern of RTs/MTs and the model's parameters (in the vein of 414 Schmitz & Voss, 2014). 415

416 Distinct inhibitory mechanisms

We introduced and fitted an algorithmic model of motor imagery allowing to 417 disentangle distinct underlying inhibitory mechanisms. Results from this analysis showed 418 that motor inhibition may have different effects on patterns of RTs and MTs, suggesting 419 that it may affect different aspects of the latent activation function (i.e., the shape of the 420 activation function or the motor execution threshold). More precisely, the shortening of 421 RTs jointly observed with the lengthening of MTs following imagined trials (as observed in 422 the present study) is compatible with a *decrease* in the motor execution threshold (as 423 reported in Table 3), whereas the lengthening of both RTs and MTs following imagined 424 trials (as observed in previous studies such as Bart et al., 2020, 2021a, 2021b; Rieger et al., 425 2017) is compatible with a delay in the peak time of the activation function (as previously 426 suggested by Nalborczyk et al., 2024). 427

As mentioned previously, task requirements during the RT period in the present 428 study differed considerably from that of previous studies. These differences in task 429 requirements may have impacted the patterns of RTs/MTs as well as the model's 430 estimates. We further speculate that across typing and pointing, motor inhibition resulting 431 from imagery affects different aspects of the execution/imagery initiation in the following 432 trial. For instance, Haith et al. (2016) showed that motor preparation (i.e., preparing a 433 movement) and motor initiation (i.e., initiating a movement once it is prepared) may be 434 modulated independently and suggested that these two components have distinctly neural 435 correlates. We therefore speculate that these two components of the RT (i.e., preparation 436 + initiation) may be impacted differently in typing vs. pointing movements and may be 437 mapped to distinct parameters of the algorithmic model. More precisely, because motor 438 preparation is associated with activity in motor and premotor cortices and may share 439 resources with motor execution, we might expect it to impact the shape of the activation 440 function – this would correspond to modulations of the peak time and/or the curvature 441 parameter. In contrast, motor initiation is not expected to affect the shape of the 442 activation function and may therefore be related to modulations of the threshold parameter 443 (whose modulations do not change the shape of the activation function), with correlates in 444 brain regions such as the SMA (as suggested for motor initiation by Haith et al., 2016). 445 Relating this discussion to the two tasks mentioned previously (typing vs. pointing), we 446 could speculate that motor inhibition in the pointing action-mode switching tasks impacts 447 the shape of the activation function (thus, slowing down motor preparation) whereas motor 448 inhibition in the typing action mode switching tasks would impact the motor execution 449 threshold (thus, slowing down or delaying motor initiation). This remains to be assessed in 450 future studies combining behavioural measures with neuroimaging or electrophysiological 451 measures of motor preparation and motor imagery. 452

Perhaps surprisingly, fitting the algorithmic model of motor imagery presented in
Nalborczyk et al. (2024) to our data revealed that the motor execution threshold was

decreased in imagined-imagined sequences relative to executed-imagined sequences. 455 Whereas this interpretation may seem unintuitive at first, it reflects the *increase* of the 456 motor execution threshold in the first (previous) imagined trial, which then drifts back to 457 an average value. In other words, the motor execution threshold decreases in the second 458 trial of an imagined-imagined sequence (relative to the first trial) because it has increased 459 in the first trial of this sequence (relative to the trial preceding this sequence) and has no 460 reason to increase further. This could reflect strategic modulations of the execution 461 threshold in sequences of trials wherein the threshold is modulated according to task 462 demands in order to minimise a putative cost associated with upregulating the threshold 463 higher than necessary. This hypothesis is supported by considering what happens in 464 sequences of three successive trials (i.e., imagined-imagined-imagined 465 vs. executed-imagined-imagined sequences). In such sequences, the motor execution 466 threshold in the last trial of this sequence is *lower* in imagined-imagined than in 467 executed-imagined-imagined sequences (RTs/MTs are smaller/longer in III than EII 468 sequences, cf. Figure S3 in supplementary materials). 469

These inhibitory mechanisms may be subserved by brain regions such as the 470 pre-supplementary motor area (Kasess et al., 2008) or the right inferior frontal gyrus, 471 which may plausibly be responsible for weakening the motor commands that are emitted 472 during motor imagery (e.g., Angelini et al., 2015, 2016; Nalborczyk et al., 2022). 473 Downstream regions in the cerebellum (e.g., Lotze et al., 1999), in the brainstem (e.g., 474 Jeannerod, 2001, 2006), or at the spinal level (e.g., Grosprêtre et al., 2016) may also 475 contribute to motor inhibition at a later stage. Alternatively, the need for explicit 476 inhibitory mechanisms counteracting motor commands has been questioned by studies 477 looking at population-level neuronal activity during motor preparation or motor imagery 478 (e.g., Dekleva et al., 2024; Kaufman et al., 2014). Analyses of intracranial recordings in 479 monkeys have revealed that neuronal activity within the primary motor cortex cancels out 480 at the population level during motor preparation, explaining how the motor cortex may 481

prepare movement without executing it (Kaufman et al., 2014). Dekleva et al. (2024) 482 further showed, in a implanted human tetraplegic patient, that motor imagery and motor 483 execution do not occupy completely orthogonal neural subspaces. Rather, population-level 484 neural activity can be decomposed in a common subspace plus orthogonal subspaces for 485 executed and imagined movements. This population-level view does not necessarily 486 contradict the algorithmic perspective of the model fitted on the present data, however. 487 Indeed, premotor areas (such as the dorsal premotor cortex) may be responsible for sending 488 "trigger" signals responsible for switching from the execution or imagery neural subspaces 480 (similar to what has been hypothesised for motor preparation and execution, e.g., Zimnik 490 & Churchland, 2021), with similar behavioural consequences (e.g., on imagined RTs and 491 MTs) than the threshold modulation perspective. Moreover, the "activation timecourse" 492 assumed by the algorithmic model may correspond to the population-level aggregate 493 activity within the primary motor cortex. Whether this assumption holds (or whether the 494 activation timecourse may be related to fluctuations in the vividness of motor imagery 495 percepts) remains to be assessed in future studies using TMS or intracranial measurements 496 together with continuous subjective reports. 497

498 Conclusions

In summary, we adapted the action mode switching paradigm (Rieger et al., 2017) to 490 typing imagery and observed that typing imagery slows down subsequent imagined or 500 executed typing movements. We further used and discussed a novel algorithmic model of 501 motor imagery allowing to infer the underlying continuous timecourse of motor activation 502 throughout the trial and to distinguish between different potential inhibitory mechanisms. 503 The assumptions and novel predictions derived from this framework remain to be assessed, 504 thus providing promising avenues for future research on the rich subjective experience of 505 motor imagery and the neural mechanisms supporting it. 506

507

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Data availability statement

⁵¹⁹ Data and supplementary materials are available via the Open Science Framework: ⁵²⁰ https://osf.io/y9a3k/.

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Figure 2. Average reaction time (top), movement time (middle), and total response time (bottom) across conditions. The error bars represent the 25th and 75th percentiles. Each dot represents the by-participant median reaction time (top row) or movement time (bottom) computed across 96 trials per condition. Left panel: imagined trials. Right panel: executed trials. Sequences of trials (i.e., pair of successive trials) are represented on the x-axis, which corresponds to (from left to right): imagined-imagined, executed-imagined, imagined-executed, and executed-executed sequences.



Figure 3. Estimated mean reaction time (top) and movement time (bottom) across conditions in Experiment 1. The error bars represent the 95% credible intervals. Sequences of trials (i.e., pair of successive trials) are represented on the x-axis, which corresponds to (from left to right): imagined-imagined, executed-imagined, imagined-executed, and executed-executed sequences.



Figure 4. Relation between the average movement time in executed and imagined trials. Each panel represents a participant, each dot/triangle represents the per-word average MT in executed and imagined trials for this participant. Dashed lines represent the isochrony lines for which executed and imagined MTs are equal. Participants classified as 'off-task' according to this correlation are depicted in orange on the left whereas participants classified as 'on-task' according to the correlation between executed and imagined MTs are depicted in green on the right.



Figure 5. Average (median) reaction time and movement time per condition for participants with a null correlation between average executed and imagined movement times (left panel), those with a weak positive correlation (middle panel), and those with a strong positive correlation (right panel).



Figure 6. Average latent activation function as estimated by the model in imagined-imagined (blue) and executed-imagined (orange) sequences from stimulus onset (time = 0) to RT (first threshold crossing) onto to MT (difference between the second and first threshold crossing). The horizontal dotted lines represent the motor execution thresholds whereas the dashed horizontal lines represent the motor imagery thresholds. Only participants with a strong positive correlation between executed and imagined MTs (N = 16) were included.