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Adjusted trajectory of medication exposure taking into account the periodicity of dispensations and the number of dispensed packs and comparative analysis on EFEMERIS database

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Abstract

We presented an adjustment method for the calculation of medication exposure trajectories based on the number of dispensed packs and the type of dispensations (occasional or regular). A comparative study based on the EFEMERIS data was carried out using three different scenarios of trajectory calculation depending on whether or not the number of packs and the periodicity of medication dispensations were taken into account. The impact of the scenario was highlighted using global indicators on the number of Define-Daily Dose on all women exposed; the study of changes in individual trajectories from one scenario to another was carried out; we also compared the results of a clustering into four groups. If 65% of the trajectories remained unchanged, we could observe on the rest significant changes in number of DDD and/or on individual exposure profile. We observed 4% of trajectories that were attributed to a different cluster, and the clustering was of better quality with the adjustment method. Depending on the study context, an impact on cluster distribution could be observed for some maternal characteristics and neonatal outcomes. This was the case for a higher occurrence of neonatal pathology for neonates from mothers belonging to the cluster with high doses of psychotropics, thus reinforcing the conclusions of previous studies of a link between high exposure to psychotropic medications and presence of pathology for the newborn.

Introduction the set of the set of

The challenge in measuring medication exposure lies in the different types of information available in health databases like dispensed medications, dispensing dates, ³ number of dispensed packs, number of units per pack. A minimalist approach, and ⁴ widely used so far, is to transform this information into exposed/non-exposed (possibly $\frac{1}{5}$ over time) $[1, 2]$. This binary approach does not consider important aspects such as

dosage, treatment duration, and timing of exposure, and leads to the quantitative and ⁷ temporal aspect of the medication load being missing. In a previous study $[3]$, a new \bullet pharmacoepidemiological method has been developed to take into account intensity and ⁹ evolution of medication exposure and to estimate medication load in the form of an 10 exposure trajectory over time. This method proposed in $[3]$ has been applied to $\frac{11}{11}$ pregnant women exposed to psychotropic medications from the EFEMERIS database, a ¹² cohort of pregnant women and their outcomes in South-West France $[4, 5]$. We proposed $_{13}$ to expand this method using another mode of calculating the medication load, which $_{14}$ takes into account both the periodicity of medication dispensations and the number of $_{15}$ dispensed packs.

We first recall the initial method and then describe the one developed. To assess the $\frac{1}{10}$ numerical contribution of the new calculation method versus the initial method, a comparative analysis was performed based on three scenarios, depending on whether or ¹⁹ not the periodicity and the number of dispensed packs were taken into account. The 20 main changes due to the proposed method on the medication exposure distribution and ₂₁ on exposure clusters are detailed in Section Results. ²²

Materials and methods 23

Initial method of calculating exposure trajectories 24

To quantify exposure to all dispensed medications, physical quantities of medications in ²⁵ number of units were transformed into a standard unit of measurement proposed by the ₂₆ World Health Organisation: the Defined-Daily Dose, denoted by DDD [6]. Women 27 taking psychotropic medications are often exposed to several ones (for example, ²⁸ antidepressant and anxiolytics) and the use of DDDs allows to aggregate units of ²⁹ different medications (in our case, the different types of psychotropic medications). Exposure periods were computed from dispensation dates and number of units per pack. 31 The number of dispensed packs was not available and set at 1 by default. In a general $\frac{32}{2}$ manner, it is assumed that for each dispensed medication, women are exposed to one $\frac{33}{2}$ DDD per day (cf. $[3]$). For example, for one dispensed pack of 7 DDD, the treatment $\frac{34}{4}$ duration was estimated at 7 days with one DDD per day. If 14 DDDs were dispensed, $\frac{35}{25}$ the treatment duration was estimated at 14 days. Thus, in the initial mode of $\frac{36}{10}$ calculation, the treatment periods of the same active substance were placed one after $\frac{37}{20}$ the other. In case of pregnancy, some DDDs could be shifted after birth and were $\frac{38}{10}$ therefore not taken into account in the calculation of the exposure measurement. This ³⁹ computation could be apply to a medication, or several medication of the same class: ⁴⁰ after superimposing individual periods of exposure, the number of DDDs per day and ⁴¹ per medication was cumulated for all psychotropic medications dispensed on a given day ⁴² of a woman's pregnancy. A sequence of daily quantitative measurements of psychotropic ⁴³ exposure was therefore available for each woman, which could be called daily exposure ⁴⁴ trajectory.

Proposed method of calculating exposure trajectories based on $\frac{46}{46}$ periodicity adjustment 47

In case where the number of dispensed packs was available, the first step was to ⁴⁸ calculate the total number of units dispensed at each dispensation of a specific ⁴⁹ medication (identified using the ATC code) in order to estimate the active medication \sim load closer to reality. Then, the periodicity was included by differentiating single or $\frac{51}{10}$ occasional dispensing from regular dispensing. A dispensation was considered as single $\frac{52}{2}$ or occasional if it was not renewed over a period of 45 days. In this case, the $\frac{53}{100}$

distribution of the corresponding total DDDs was done as for to the initial method, ie. $\frac{54}{100}$ put one after the other with one DDD per day. Dispensations were defined as regular if $\frac{55}{15}$ they were repeated in a period of less than 45 days, based on French dispensing habits $\frac{56}{10}$ for one month of treatment, adding a flexibility of 15 days. If a specific medication was 57 dispensed regularly for a woman, we calculated the average number of DDDs dispensed ss over all regular dispensations (\overline{N}_{DDD}) and the average duration between two regular $\frac{59}{2}$ dispensations $(\overline{D}_{between})$. We deduced the ratio $F = \overline{N}_{DDD}/\overline{D}_{between}$ which estimated, 60 for the considered pregnancy, the number of DDDs per day distributed between the first ϵ_1 and the last regular dispensation. From the last dispensation, the remaining number of ϵ DDDs per day was $F/2$ until exhausted to consider possible weaning.

As the initial method, we then deduced the total number of DDDs per day by cumulating all medications of interest (psychotropic) dispensed on a given day of a 65 woman's pregnancy. In this study, we also cumulated the exposure measurement by $\frac{66}{66}$ week, to obtain weekly exposure trajectory for each exposed pregnancy.

\mathbf{Data} 68

Data for our analysis comes from the EFEMERIS database $[4, 5]$. The EFEMERIS 69 cohort was approved by the French Data Protection Authority on 7 April 2005 (authorization number 05-1140). This study was performed on anonymized patient data. τ_1 The women included in the EFEMERIS database were informed of their inclusion and $\frac{72}{2}$ of the potential use of their anonymized data for research purposes. They could oppose $\frac{73}{2}$ the use of their data at any time. The women included in the EFEMERIS database know $_{74}$ that their collected and anonymized data can be used for medical research purposes and $\frac{75}{75}$ can thus be published. The study was approved by the EFEMERIS steering group. $\frac{76}{6}$ Data were handled and stored in accordance with the General Data Protection $\overline{77}$ Regulation. The EFEMERIS database contains data about medication dispensed to $\frac{78}{8}$ pregnant women (from CPAM, the national medical insurance organisation) in ⁷⁹ Haute-Garonne, data concerning pregnancy outcomes and data on the children obtained \bullet from mother-and-child health services (PMI, "Protection Maternelle et Infantile", which ⁸¹ records information about the child and delivers the mandatory health certificates for $\frac{1}{82}$ children at eight days, nine months and 24 months of age). We also have information on $\frac{83}{100}$ pregnant women such as age, presence of diabetes or preeclampsia. The present study ⁸⁴ was restricted to mother-child pairs included in EFEMERIS between June 1, 2011 and ⁸⁵ December 31, 2020. Over this period, we retrieved $24,138$ psychotropic medication $\frac{86}{100}$ dispensations among 6,820 live pregnancy outcomes for which the presence or absence of $\frac{87}{87}$ neonatal pathology was known. The statistical unit of our study was the pair, pregnancy-outcome of pregnancy, which will be noted pregnant woman or pregnancy in ⁸⁹ the following document in the absence of ambiguity. In this context, for each pregnancy, $\frac{90}{200}$ the exposure trajectory described the exposure to psychotropic medications. More 91 precisely, the exposure trajectories started from 4 weeks before pregnancy to take into $\frac{92}{2}$ account the potential long-term presence of medications in the woman's body and its $\frac{93}{2}$ potential effect on the fetus from the start of pregnancy. To calculate such an exposure ⁹⁴ trajectory from 4 weeks before pregnancy, medication dispensations from week -13 were $\frac{1}{95}$ used, because a treatment taken at -4 have been dispensed between -13 and -4. But 96 DDD before -4 were not considered for analysis of exposure trajectories. Weekly $\frac{97}{97}$ measures of exposure to psychotropic medications were calculated for each pregnant woman exposed at least once between -4 weeks before pregnancy and the end of 99 pregnancy, called pregnancy period in the rest of this document. ¹⁰⁰

Remember that the trajectories were calculated from dispensing from 13 weeks before 130 pregnancy. This could impact whether a pregnancy was classified as exposed from 4 ¹³¹ weeks before pregnancy and thus could change the exposure trajectory. Taking into account the number of dispensed units (from scenario 1 to 2) allowed to recover 94 133 pregnancies, and adding the periodicity adjustment (from scenario 2 to 3) allowed to ¹³⁴ recover 29 more pregnancies, ie. an increase of 2.4% in the number of pregnancies 1_{35} exposed at least one week between -4 and 38 weeks of pregnancy (second column of ¹³⁶ Table 1). The sum of DDDs increased by 32% from scenario 1 to 2 and by 42% from 137 scenario 1 to 3 across global trajectories of all exposed pregnancies (third column of $_{138}$ Table 1). Considering only the impact of the periodicity adjustment (from scenario 2 to 139 3), the total DDDs increased globally by 7% .

Note that the increase in the sum of DDD between scenarios 1 and 2 was due to the ¹⁴¹ fact that there were more pregnancies classified as exposed and that exposure measures $_{142}$ were logically increased if the number of units is greater than 1. The increase from $\frac{143}{143}$ scenario 2 to 3 was also explained by the fact that in scenario 2, some DDDs were $_{144}$ shifted after the end of pregnancy, while with scenario 3, the periodicity adjustment 145

allowed them to be positioned more appropriately during pregnancy period. The ¹⁴⁶ average number of weekly DDD by pregnancy increased for example from 47.3 for $_{147}$ scenario 1 to 65.6 for scenario 3, ie. almost 20 DDD more over the entire pregnancy $_{148}$ period (from -4 to around 38 weeks). We also studied the weekly DDD distribution ¹⁴⁹ (four last columns of Table 1). The periodicity adjustment had the effect of largely ¹⁵⁰ increasing the maximum DDDs as well as the largest values (see the 95 percentiles). $\frac{151}{151}$ But the average increased slightly, due the fact that at least 75% of the DDDs were $_{152}$ equal to 0 for each of the three scenarios.

	Number of exposed pregnancies	Sum of DDDs between -4 and 38	Distribution of weekly DDD			
	at least one week			90	-95	
	between -4 and 38		Mean (sd)	percentile	percentile	max
Scenario 1	5468	258427	1.13(3.00)	6.5	7.0	47
Scenario 2	5562	341656	1.47(3.60)	7.0	7.0	50
Scenario 3	5591	366485	1.57(4.06)	6.5	8.7	84

Table 1. Indicators of comparison of DDD values between the three scenarios

$Step 2: Impact on individual trajectories$

Step 2 considered the pairwise differences between the three trajectories of each exposed 155 pregnancy. From scenario 1 to 2 (adding the number of dispensed units), 68% of 156 pregnancies had a different trajectory, ie. at least one week with a different DDD, which ¹⁵⁷ was consistent with the percentage of dispenses of only one pack. Only 10% of 158 trajectories increased on average by at least one DDD per week, and up to 10 DDD per 159 week. 160

From scenario 2 to 3 (adding the periodicity adjustment), 65% of the trajectories $_{161}$ were strictly identical. Among the 35% of pregnancies having different trajectories, one ¹⁶² third had a total of DDDs decreased with a maximum drop corresponding to an average 163 of 3 DDDs per week. This decrease could be explained mainly by the modified ¹⁶⁴ calculation of DDD before -4 weeks which could lead to a greater number of DDDs 165 before week -4 and therefore lower after. Half of the modified trajectories realized an ¹⁶⁶ increase of the total DDDs corresponding to 1.3 DDDs per week on average with $_{167}$ standard deviation equal to 2.1. For the rest of these modified trajectories, the total of $_{168}$ DDDs did not change due to a compensation effect between weeks. We observed 2% of $_{169}$ all trajectories (ie. 106 trajectories) associated with an augmentation of the total of 170 DDD of between approximately 2 and 20 DDD per week on average.

Fig 1 represents examples of trajectories calculated according to the 3 scenarios for 4 μ exposed pregnancies. For these four pregnancies, the number of dispensations was between 24 and 39 corresponding to between 2 and 7 different dispensed ATC, with half ¹⁷⁴ of dispensations of more than one pack. In the three first examples, we observed that ¹⁷⁵ taking into account the number of packs was mainly enhanced by the periodicity ¹⁷⁶ adjustment. This revealed changes in the trajectory profiles: a higher exposure at the $\frac{177}{20}$ start of pregnancy, followed by a gradual decrease for pregnancy $n^{\circ}1$, a higher exposure $\frac{178}{178}$ throughout the pregnancy period for pregnancy n^2 and a medication load multiplied by $\frac{179}{2}$ 2 after the first trimester of pregnancy for pregnancy n[°]3. The changes were less 180 significant for the pregnancy n^2 .

$Step 3: Impact on clustering$ 182

Using the individual weekly exposure trajectories calculated with scenario 2 and 3, we 183 used the Kmeans method for longitudinal data to create exposure clusters. We 184

 $\frac{5}{8}$

Fig 1. Examples of individual exposure trajectories. Examples of exposure trajectories calculated according to three calculation scenarios, for four exposed pregnancies. The woman of pregnancy n°1 received 24 medication dispensations of 3 different ATC over the pregnancy period, of which 20 with more than one pack. The woman of pregnancy n°2 received 36 medication dispensations of 2 different ATC over the pregnancy period, of which 32 with more than one pack. The woman of pregnancy n°3 received 25 medication dispensations of 2 different ATC over the pregnancy period, of which 13 with more than one pack. The woman of pregnancy n°4 received 39 medication dispensations of 7 different ATC over the pregnancy period, of which 19 with more than one pack.

evaluated the impact of using the periodicity adjustment on the creation of exposure 185 clusters by comparing the clustering from scenario 2 and 3. The quality of the ¹⁸⁶ clustering measured by the Calinski-Harabatz criterion was better for scenario 3: 2302 ¹⁸⁷ for scenario 3 versus 2010 for scenario 2. Based on four clusters (see [3]), it could be $_{188}$ seen that taking into account the periodicity adjustment did not change the ¹⁸⁹ representative trajectory profile of the four clusters (see Fig 2). The average level of 190 exposure by cluster remained stable, except for the cluster corresponding to the highest ¹⁹¹ exposures: on average, over individual trajectories, 706 DDD (standard deviation=291) $_{192}$ for scenario 2 versus 800 (sd=267) for scenario 3). The study of the concordance between the two partitions showed that 4% of trajectories of exposed pregnancies $\frac{194}{2}$ changed clusters between the two scenarios.

We studied the impact of these 4% of trajectories on the cluster distribution of some $_{196}$ maternal characteristics (such as pre-eclampsia and diabetes) and neonatal outcomes 197 (such as neonatal pathologies or preterm birth) not taken into account in the formation ¹⁹⁸ of the clusters. Concerning variables affecting the mother, it was observed for example ¹⁹⁹ that the proportion of pre-eclampsia per cluster remains unchanged from one scenario $_{200}$ to another, but there is a higher proportion of diabetes reported among pregnancies in $_{201}$ the high-exposure cluster in scenario 2 (19.5%) as compared with scenario 1 (15%). For $_{202}$ variables concerning the newborn, the proportion of premature births remained 203 unchanged from one scenario to another. But it was observed that the proportion of \qquad 204 neonatal pathology in high-exposure clusters increased with the periodicity adjustment ²⁰⁵ from 21.6% (out of 111 pregnancies) to 26% (out of 92 pregnancies).

These elements suggested that depending on the study conducted and the choice of $_{207}$ considered variables, taking into account the periodicity adjustment does not change the ²⁰⁸ conclusions of the study, but can refine them. In the case of the presence of neonatal ²⁰⁹

pathology, it reinforced the results of the study conducted in $[7]$ on the link between $\qquad 210$ exposure to psychotropic medications and the presence of neonatal pathology. On this $_{211}$ example with the EFEMERIS database, taking into account or not of the periodicity 212 adjustment did not radically modify the results, but the quality of the clustering is 213 better and the trend of the main results is strengthened.

Fig 2. Individual exposure trajectories and mean trajectories for clusters. Four clusters were identified by K-means method for longitudinal data. Individual trajectories in the same cluster are plotted in the same color. The top figure represents the exposure trajectories calculated according to scenario 2, and the bottom one according to scenario 3.

Conclusion

In the case where the number of dispensed packs was available, we highlighted an impact $_{216}$ of adjusting the calculation of exposure measure by taking into account the periodicity ²¹⁷ of dispensations. The analyzes on the total number of DDDs over all trajectories or on ²¹⁸ individual ones showed that the method increased the average exposure measure quite ²¹⁹ slightly. 65% of the exposed pregnancies did not have any change in the weekly trajectories using the different methods. Among the changed trajectories, 90% showed $_{221}$ an increase or a decrease of the cumulative medication exposure, and the remaining 10% 222 had changes according to the weeks that are compensated. There could be a great 223 variability in changes in the medication exposure and thus for a non-negligible number $_{224}$ of trajectories, an increased number of DDDs and a modified exposure trajectory profile ²²⁵ could be observed. The adjusted trajectories led to clusters with no radical ²²⁶ modifications but with better quality and higher level of exposure for the women most 227 exposed. Considering others variables, adjustment variables or outcomes in ²²⁸ pharmacoepidemiology studies, their distribution by cluster could be impacted or not. ²²⁹

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