

# Empirical Orthogonal Functions for ecology

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## ▶ To cite this version:

Baptiste Alglave, Benjamin Dufée, Said Obakrim, James T Thorson. Empirical Orthogonal Functions for ecology. 2025. hal-04693871v2

## HAL Id: hal-04693871 https://hal.science/hal-04693871v2

Preprint submitted on 11 Feb 2025

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## 1 Journal: Ecology

## 2 Submission Type: Statistical Report

## 3 *Title:* Empirical Orthogonal Functions for Ecology

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- 6 Keywords: spatio-temporal data, dimension reduction, satellite data, NDVI,
- 7 species distribution model, multivariate analysis, canonical correlation
- 8 analysis.

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Abstract: Spatio-temporal data are ubiquitous in ecology, involving large 9 spatial domains, long time periods, and multiple variables. To take 10 advantage of these data for ecological analysis, it is crucial to rely on both 11 efficient and ecologically interpretable statistical methods. Empirical 12 Orthogonal Functions (EOFs) are a widely used method for reducing the 13 dimensionality of spatio-temporal data. While EOFs are common in 14 meteorology and climate studies, their use in ecology is still emerging. 15 This paper introduces the basics of EOFs and presents new methods to 16 specifically address key ecological questions. Since EOFs lack explicit 17 ecological constraints, their results may be difficult to interpret. We propose 18 a method to rotate EOFs using a temporal ecological variable, improving 19 interpretability by better representing ecological patterns. We also explore 20 the use of EOFs in multivariate analysis to identify shared spatial and 21

22 temporal patterns in community and ecosystem dynamics.

We illustrate these methods on two case studies including NDVI satellite
data over France and spatio-temporal predictions from fish species
distribution models.

Our goal is to provide a clear and concise roadmap for applying EOFs to
ecological problems. We emphasize practical guidance through code
snippets and two worked examples, making the approach accessible to a
broad audience of ecologists.

#### 30 1 Introduction

Spatial and temporal data are ubiquitous in ecological research. Survey data, citizen-science data, satellite data or biogeochemical model outputs provide huge information for unraveling ecological processes at various temporal and spatial scales (Hampton et al., 2013; Grémillet, Chevallier, and Guinet, 2022). Given the extensive spatial and temporal coverage of all these datasets, dimension reduction (or ordination) methods become critical for synthesizing the information embedded within these spatio-temporal data.

Empirical Orthogonal Functions (EOFs) represent the keystone method for 38 performing dimension reduction on spatio-temporal data. EOFs were first 39 introduced by Lorenz (1956). Basically, they are Principal Component 40 Analysis (PCA) performed on spatio-temporal data. Since their first 41 formulation, EOFs methods have entailed a rich literature, particularly in 42 meteorology and climate applications (Hannachi, Jolliffe, and Stephenson, 43 2007). Several extensive overviews of EOFs and derived methods are 44 available in the secondary literature (Preisendorfer, 1988; Von Storch and 45 Zwiers, 2002; Cressie and Wikle, 2011; Hannachi, 2021); they mainly 46 address an audience of climatologists or statisticians. 47

Recently, these methods have expanded to ecology. They have been applied
to biogeochemical variables (Schrum, John, and Alekseeva, 2006; Iida and
Saitoh, 2007; Di Lorenzo et al., 2008; Woillez et al., 2010; Le Mezo et al.,

2016; Berkelhammer, 2019) as well as species distribution (Norton and 51 Mason, 2005; Petitgas, Masse, et al., 2006; Petitgas, 2008; Woillez et al., 52 2010; Petitgas, Doray, et al., 2014; Petitgas, Huret, et al., 2018; Petitgas, 53 Renard, et al., 2020; Grandremy et al., 2023; Van Audenhaege et al., 2022). 54 EOFs have gained wider recognition through the development of the 55 Vector-Autoregressive Spatio-temporal (VAST) package (Thorson, 2019; 56 Thorson, Anderson, et al., 2024) and the application of EOFs to represent 57 the spatio-temporal variability of species communities and ecosystems 58 (Thorson, Ciannelli, and Litzow, 2020; Thorson, Cheng, et al., 2020; 59 Thorson, Arimitsu, et al., 2021). Still, these applications remain mainly 60 focused on marine systems, with very few applications to terrestrial systems 61 (Gedalof, Peterson, and Mantua, 2005; Wu et al., 2023). 62

Overall, EOFs methods can address a large range of ecological questions 63 from analyzing single species spatio-temporal variability to aiding 64 conservation decisions in an ecosystem perspectives. They allow to detect 65 seasonal signals in species distribution and shifts in phenology (Alglave, 66 Olmos, et al., 2024). Additionally, EOFs are particularly useful in tracking 67 long-term ecological changes, such as ecosystem restructuring or shifts in 68 species composition in response to climate change (Thorson, Arimitsu, 69 et al., 2021; Badger, Large, and Thorson, 2023). 70

71 Despite their broad potential, EOFs remain underutilized in ecology,

particularly in terrestrial ecosystems. In this paper, to widen EOFs 72 applications and promote their use in ecology, we provide a thorough and 73 accessible introduction to EOFs. Additionally, we describe two derived 74 methods to address specific ecological questions: (1) a method that sets 75 ecological constraints on EOFs to better disentangle the information of the 76 different dimensions of the EOFs and improve their interpretability, and (2) 77 a multivariate version of EOFs that enables to apply EOFs methodology on 78 species communities or whole ecosystems. 79

We illustrate these methods based on two real applications. First, we use a 80 terrestrial ecology example based on monthly Normalized Difference 81 Vegetation Index (NDVI) satellite data (MODIS) from 2000 to 2023 82 (MODIS, 2021). As a second case study, we use spatio-temporal predictions 83 from an Integrated Spatio-Temporal Species Distribution Model (IST-SDM) 84 (Alglave, Vermard, et al., 2023). We analyze the spatio-temporal predictions 85 spanning from 2008 to 2018 for three demersal species in the Bay of Biscay: 86 common sole (Solea solea, Linnaeus, 1758), squid (Loligo vulgaris, 87 Lamarck, 1798), and whiting (Merlangius merlangus, Linnaeus, 1758). In 88 the case of the IST-SDM, when the method is univariate (*i.e.* single species) 89 we focus the analysis on common sole. 90

Our goal is to provide a clear and concise roadmap for applying EOFs to
ecological problems. We emphasize practical guidance through code

snippets and two worked examples, making the approach accessible to abroad audience of ecologists.

## 95 2 Basics of EOFs

EOFs were initially introduced by Lorenz (1956) for the analysis and
prediction of weather. The method involves extracting the main spatial and
temporal patterns that (1) capture the most variance of spatio-temporal data
and (2) are orthogonal to each other. In the upcoming subsections, we
introduce the notations and outline the principal properties of EOFs through
both theoretical explanations and practical applications.

#### 102 2.1 Continuous formulation of EOFs

We denote a space-time process *S* such that S(x, t) represents the value of the process *S* at spatial location  $x \in \mathcal{D}$  (with  $\mathcal{D}$  a continuous domain) and discrete time step denoted *t*. When dealing with EOFs, we assume that the spatio-temporal field can be decomposed into a (purely) spatial term and a (purely) temporal term. In its continuous formulation, this decomposition can be expressed as:

$$S(x,t) = \sum_{m=1}^{\infty} \alpha_m(t) \cdot \phi_m(x)$$
(1)

where  $\phi_m(x)$  denotes the spatial term of the EOFs for dimension *m* and  $\alpha_m(t)$  represents the temporal term of the EOFs for dimension *m*. Furthermore, var  $(\alpha_1(t)) >$ var  $(\alpha_2(t)) > \dots$ , and cov  $(\alpha_k(t), \alpha_m(t)) = 0$ , for all  $k \neq m$ . A known solution of this problem is obtained through a Karhunen-Loève expansion. Both formulation and practical estimation of the EOFs in the continous case are detailed in Cressie and Wikle (2011) (p.266).

#### 116 2.2 Discrete formulation of EOFs

The most common and straightforward way to perform EOFs on *S* is to consider a discretized version of this process. Let's introduce the matrix **S** of size  $n \times p$ , where *n* denotes the number of discrete spatial positions, and *p* denotes the number of time steps, with the time stepping denoted by  $(t_1, ..., t_p)$ .

In practice, to decompose S through EOFs, these data are centered and thendecomposed through Singular Value Decomposition (SVD).

From these data, for a given spatial location x, the temporal average can be calculated as:

$$\overline{S}^{t}(x) = \frac{1}{p} \sum_{k=1}^{p} S(x, t_{k}).$$
(2)

126 The time-centered space-time field, also referred to as the anomaly matrix,127 is then obtained by:

$$\mathbf{S}' = \mathbf{S} - \overline{\mathbf{S}}^t \, \mathbf{1}^{\mathsf{T}},\tag{3}$$

where  $\overline{\mathbf{S}}^{t}$  is an  $n \times 1$  column vector of row means  $\overline{\mathbf{S}}^{t}(x)$  and  $\mathbf{1}^{\top}$  is a  $1 \times p$ row vector of ones. Codes to compute the anomaly matrix are available in Supplementary Material (SM S1.1). Examples of both the raw spatio-temporal data and anomalies are available in Figure 1 for satellite NDVI data and in Figure S3.1 (SM3) for the IST-SDM case study.

133 In this case, the decomposition can be written as:

$$S'(x,t) = \sum_{m=1}^{r} \lambda_m \cdot u_m(x) \cdot v_m(t)$$
(4)

134 Here, r represents the number of dimensions of the EOFs with

135  $r = \min(n, p)$  being the rank of the matrix **S**'.  $u_m(x)$  is a spatial term and 136  $v_m(t)$  is a temporal term; both  $\mathbf{u}_m$  and  $\mathbf{v}_m$  have unit norm.  $\lambda_m$  quantifies the 137 amplitude of the  $m^t h$  dimension of the EOFs. The  $\lambda_m$  are called the singular 138 values hereafter (see section 2.2.1).

139 The key aspect of EOFs analysis lies in the constraints imposed on the

spatial terms  $u_m(x)$  or the temporal terms  $v_m(t)$  because they are crucial for

141 the computation and interpretation of these indices. The most

straightforward and natural constraints involve (1) maximising the variance explained by each spatial pattern  $p_m(x)$ ; (2) setting orthogonal constraints for both the spatial and temporal terms.

#### 145 2.2.1 EOFs as a solution of Singular Value Decomposition

146 To obtain the terms  $u_m(x)$  and  $v_m(t)$  of equation (4), it is possible to

147 decompose S' through Singular Value Decomposition (SVD), namely:

$$\mathbf{S}' = \mathbf{U} \mathbf{\Sigma} \mathbf{V}^T \tag{5}$$

with U a  $n \times r$  matrix and V<sup>T</sup> a  $r \times p$  matrix,  $r = \min(n, p)$  being the rank 148 of the matrix S'.  $\Sigma$  is a  $r \times r$  diagonal matrix with non-increasing positive 149 coefficients on the diagonal, denoted as  $\Sigma = \text{Diag}(\lambda_1, ..., \lambda_r)$  and 150  $\lambda_1 \ge \lambda_2 \ge ... \ge \lambda_r > 0$ . This is essentially the same computation that is 151 performed when conducting PCA on any type of multivariate data. 152 153 The columns of  $\mathbf{U} = (\mathbf{u}_1, ..., \mathbf{u}_r)$  are referred to as the EOFs maps of the anomaly matrix S'. In the following, they will be referred to as factors; here 154 the factors are spatial. The columns of  $\mathbf{V} = (\mathbf{v}_1, \cdots, \mathbf{v}_r)$  contain the terms 155

156  $v_m(t)$  from Equation (4). They will be called **loadings**; here the loadings are

- 157 temporal. Finally,  $\Sigma$  contains the singular values of S' along the diagonal.
- 158 Note that here, singular values cannot be related to variance because the

centering of S is realized on rows. Codes to perform the svd are available inSM1 (Section S1.2).

#### 161 2.2.2 Interpretation of EOFs

The construction of EOFs discussed in the previous subsection implies that the first EOFs map  $\mathbf{u}_1$  is the factor that captures the most variance in **S**'. The second EOFs,  $\mathbf{u}_2$  is the next factor that captures most variance while being orthogonal to  $\mathbf{u}_1$ . This applies to the next  $j^{th}$  factors.

166 The loadings contained in  $\mathbf{v}_i$ , where  $j \in 1, \dots, r$ , are temporal indices.

167 Each is associated with a corresponding factor  $\mathbf{u}_i$ . When the loading  $v_i(t)$  is

positive, the anomaly field is distributed according to the factor  $\mathbf{u}_j$ .

169 Conversely, if  $v_j(t)$  is negative, the anomaly field is distributed according to 170  $-\mathbf{u}_j$ .

171 Also it is well known that identifying EOFs dimensions with specific

172 ecological (or physical) processes should be done with care (Monahan et al.,

173 2009). Indeed, because there is no explicit ecological criteria in the

- 174 computation of EOFs, the signal related to an ecological process could be
- 175 found on several dimensions. Besides, a single dimension could gather the
- 176 signal of different ecological processes. As so, methods derived from EOFs

that integrate ecological knowledge would help to better disentangle thesignal related to some specific ecological process.

Other details to interpret EOFs are provided in SM2 (*e.g.* computation of the variance related to each EOFs dimension, number of dimensions to retain in the analysis, whether specific dimensions can be related to physical or ecological processes).

## **3** Applying EOFs to the case studies

In order to illustrate EOFs, we use two spatio-temporal datasets to serve as a case study: (1) NDVI data from MODIS Vegetation Index Products (satellite data on terrestrial ecosystems) and (2) spatio-temporal predictions from the IST-SDM for sole in the Bay of Biscay (IST-SDM). A full description of the case studies is given in SM S3.2 and S3.3. We retain the first two dimensions for the NDVI case study and the first four dimensions in the IST-SDM case study based on the angle identified in the scree plot (left, Figure S4.1).

191 Both satellite NDVI data as well as the IST-SDM case study emphasize

192 strong seasonality 2. In the case of the NDVI, the first and second

193 dimensions bring out seasonal cycles that can be related to temperature and

194 precipitation. For the IST-SDM, the two first dimensions highlight a signal

195 related to reproduction (onshore-offshore migration) and the second

dimension emphasize a signal related to northward migration. Additionalanalysis on the EOFs results are available in SM S4.2.

As mentioned previously, ecological signals are mixed between the several dimensions of the EOFs. For instance, both the first and second dimensions for the two case studies capture similar seasonal signals (Figure 2). Also, some loadings seem to mix both long term trends and seasonnal trends (see the second EOFs dimension of the IST-SDM case study). Disentangling the ecological signals from these dimensions would benefit the ecological analysis of these datasets.

#### **4** Informing EOFs with an ancillary temporal variable

As mentioned in previous sections, EOFs may not effectively disentangle 206 the ecological processes that structure the data. To address this limitation, 207 informing EOFs with a temporal variable (or ancillary variable) that is 208 ecologically meaningful can improve the interpretability of EOFs patterns. 209 This section describes the methodology for extracting EOFs patterns that 210 exhibit the highest correlation with some ancillary temporal variable. This 211 process involves projecting the spatio-temporal data onto the EOFs basis 212 and then performing a Canonical correlation analysis (CCA) between a 213 selected set of EOFs loadings and the ancillary variable. We first introduce 214

the CCA method, demonstrate how to adapt it to our specific issue, andfinally present applications of this technique.

#### 217 4.1 Basics of canonical correlation analysis

CCA was first proposed by Hotelling (1992). It is a method used to identify 218 and measure the associations among two sets of variables. Consider two 219 space-time processes  $S^{(1)}(x, t)$  and  $S^{(2)}(x, t)$  at location x and time t. CCA 220 aims at identifying pairs of spatial basis vectors,  $\mathbf{w}_1$  and  $\mathbf{w}_2$ , for the two 221 processes,  $S^{(1)}$  and  $S^{(2)}$ , that maximize the correlation between their 222 projections onto these vectors. Consider the corresponding observed 223 space-time matrices denoted  $\mathbf{S}^{(1)}$  and  $\mathbf{S}^{(2)}$  of size  $n_1 \times p$  and  $n_2 \times p$ , 224 respectively. 225

226 The CCA maximizes the following correlation coefficient  $\rho$  between 227  $\mathbf{y}_1 = \mathbf{S}^{(1)}\mathbf{w}_1$  and  $\mathbf{y}_2 = \mathbf{S}^{(2)}\mathbf{w}_2$ :

$$\rho = \frac{\mathbb{E}(\mathbf{y}_1 \mathbf{y}_2)}{\sqrt{\mathbb{E}(\mathbf{y}_1^2)\mathbb{E}(\mathbf{y}_2^2)}} = \frac{\mathbb{E}(\mathbf{w}_1^T \mathbf{S}^{(1)T} \mathbf{S}^{(2)} \mathbf{w}_2)}{\sqrt{\mathbb{E}(\mathbf{w}_1^T \mathbf{S}^{(1)T} \mathbf{S}^{(1)} \mathbf{w}_1)\mathbb{E}(\mathbf{w}_2^T \mathbf{S}^{(2)T} \mathbf{S}^{(2)} \mathbf{w}_2)}}.$$
 (6)

Here  $\mathbf{y}_1$  and  $\mathbf{y}_2$  are the temporal variables that arise from the linear combination of  $\mathbf{S}^{(1)}\mathbf{w}_1$  and  $\mathbf{S}^{(2)}\mathbf{w}_2$  and whose correlation coefficient is to be maximised. Then, the spatial basis vectors  $\mathbf{w}_1$  and  $\mathbf{w}_2$  describe which areas contribute to this correlation in the dataset  $\mathbf{S}^{(1)}$  and  $\mathbf{S}^{(2)}$  respectively. In order to find the vectors  $\mathbf{w}_1$  and  $\mathbf{w}_2$  that maximize  $\rho$ , we resort to the method of Lagrange multipliers, leading to a generalized eigenvalue problem (Weenink, 2003).

$$\begin{cases} (\mathbf{C}_{11}^{-1}\mathbf{C}_{12}\mathbf{C}_{22}^{-1}\mathbf{C}_{21} - \rho^2)\mathbf{w}_1 = 0 \\ (\mathbf{C}_{22}^{-1}\mathbf{C}_{21}\mathbf{C}_{11}^{-1}\mathbf{C}_{12} - \rho^2)\mathbf{w}_2 = 0 \end{cases}$$
(7)

Here, the eigenvalues  $\rho^2$  represent the squared canonical correlations, and the eigenvectors  $\mathbf{w}_1$  and  $\mathbf{w}_2$  are the canonical correlation basis vectors.

#### 237 4.2 Coupling CCA with EOFs

To inform the EOFs with a single ancillary variable, it is possible to perform a CCA on the matrix **V** (Equation 5) composed of the loadings and an ancillary variable that we denote  $An(t), t \in \{1, \dots, p\}$ . Usually, to apply CCA we only retain the dimensions of the EOFs that capture the process and leave apart the remaining dimensions (see the rule of thumb in SM2). In this case, **w**<sub>1</sub> are coefficients that allow to rotate the loadings as well as the related factors to obtain:

one temporal variable that closely fit the ancillary variable. This is the
 combination of the EOFs loadings that maximizes the correlation
 with the ancillary variable.

the related factor (or spatial basis vector) that is a combination of the 248 first EOFs factors. It will capture as much variance as the first pattern 249 of the EOFs but will not be orthogonal to the following EOFs. 250 Note that the two datasets  $S^{(1)}$  and  $S^{(2)}$  do not necessarily have the same 251 spatial extent, but they must have the same time stepping. 252 Such procedure will not necessarily bring out strongly different patterns 253 compared with EOFs; it will rather help to disentangle the processes 254 identified in the EOFs to ease interpretation. 255

#### 256 4.3 Illustration

In this section, we inform EOFs of the NDVI satellite data with an ancillary variable representing seasonality. The ancillary variable chosen here is a steady seasonal signal evidencing seasonality with positive peaks in summer (July) and negative peaks in winter (January - red variable in Figure 3). This variable aims at disentangling the common seasonal signal shared by the two first dimensions of the EOFs (Figure 2). The seasonal signal is parameterized as  $An(t) = A \cdot sin(f \cdot t + D)$ .

To have a ancillary variable that match seasonal cycles (opposite peaks in January and July) we set A = 1 (the amplitude of the signal),  $f = 1/12 \times 2\pi$ the frequency of the signal components and the delay component D = 1. Note that this procedure is not invariant to a change in the phase of the
seasonal signal and should be fixed with care. We then perform a CCA
between the first two loadings of the EOFs (components of the matrix V in
Section 2.2) and the seasonal signal. Codes to perform the analysis are
available in the SM section S1.3.

The loadings obtained from the CCA emphasize a very good fit to the 272 ancillary variable (R2 = 0.98, Figure 3, top right). By comparison, the first 273 two EOFs loadings have positive, but lower correlation coefficient (R2 = 274 0.82 and R2 = 0.57). The spatial basis obtained through CCA displays a 275 stronger North/South gradient of NDVI compared with the two first EOFs 276 (Figure 2). NDVI are high in orange/red areas during summer (i.e. the center 277 and mountain part of France) which corresponds to high NDVI in 278 mountains and all the southern part and eastern part of France. In these 279 areas, either ice melt and temperatures become favorable for vegetation 280 growth (mountains), either temperature are high and rains remain frequent 281 which increases vegetation coverage (eastern and south western part of 282 France). In winter, these areas are colder and they can be covered by snow 283 which do not favor vegetation growth. In blue areas, NDVI is relatively 284 higher during winter. These areas have a cooler climate during winter and 285 strong precipitation which favors vegetation growth. 286

287 A similar analysis is performed on the IST-SDM outputs. It allows to better

evidence the seasonal cycle of sole reproduction and the offshore-onshoregradients that shapes sole migration (Figure S4.3)

#### 290 5 EOFs for multivariate analysis

When studying entire ecosystems or community dynamics, multiple 291 variables need to be analyzed jointly using multivariate statistical methods 292 (Ovaskainen and Abrego, 2020; Thorson, Ciannelli, and Litzow, 2020). The 293 EOFs framework is flexible and adapts to the multivariate case. It allows to 294 identify the main modes of variability across several variables and to detect 295 shared spatial patterns and temporal trends. In such cases, the underlying 296 EOFs theory remains unchanged (Section 2); only the matrix to be 297 diagonalized S' is modified. 298

#### 299 5.1 Two alternatives for conducting multivariate EOFs

Let  $v \in \{1, \dots, s\}$  denotes the index of the different variables in the analysis, and let  $S^{(v)}(x, t)$  denotes the value of the space-time process for location *x*, time *t*, and variable *v*. We denote by  $\mathbf{S}'^{(v)}$  the  $n \times p$  matrix of the spatio-temporal variable *v*. Two options are available when conducting multivariate EOFs: 305 1. Construct the matrix by stacking the matrices row-wise. In this case, 306 the matrix is denoted as  $\mathbf{S}'^{(row)}_{multi}$  with dimensions  $(n \cdot s) \times p$ , and it has 307 the following structure:

$$\mathbf{S}_{multi}^{\prime(row)} = \begin{pmatrix} \mathbf{S}^{\prime(1)} \\ \vdots \\ \mathbf{S}^{\prime(s)} \end{pmatrix}$$
(8)

308 2. Alternatively, one can construct the matrix by stacking the matrices 309 column-wise. In this case, the matrix is denoted as  $\mathbf{S}'_{multi}^{(col)}$  with 310 dimensions  $n \times (p \cdot s)$  and is structured as follows:

$$\mathbf{S}_{multi}^{\prime(col)} = \left(\mathbf{S}^{\prime(1)} \quad ; \quad \cdots \quad ; \quad \mathbf{S}^{\prime(s)}\right) \tag{9}$$

311 Similar to standard EOFs, SVD enables to compute the factors U and the loadings V (Section 2). However, depending on the data stacking of the 312 anomaly matrix, these matrices do not have the same dimensions and do not 313 lead to the same interpretation. The first case is a temporally synthetic 314 **representation** of the multivariate data. For each dimension, there is only 315 one single time series of loadings and there are as many factors as there are 316 variables. The loadings quantify how the different factors evolve over time. 317 This allows to summarize the temporal variation of all the variables in one 318 single temporal variable. A similar approach is adopted in the VAST 319

package (Thorson, 2019). The second case, is a spatially synthetic 320 representation of the data. For each dimension, there is only one factor that 321 summarizes the variability of all the variables and as many time series of 322 loadings as there are variables. The loadings of the different variables 323 quantify how these variables are related to the factor at a specific time step. 324 Note finally that in some cases, multivariate EOFs can be applied on a 325 single variable with some time-lag. In this case the different variables are 326 the lagged version of the spatio-temporal data. This typically allows for a 327 better identification of temporal patterns such as periodicity or long-term 328 trends. It is referred to as lag EOFs in this paper; it is also often referred as 329 extended EOFs in the literature (Weare and Nasstrom, 1982). 330

#### 331 5.2 Illustration

To illustrate multivariate EOFs analysis, we apply the two alternative 332 approaches to the three species in the Bay of Biscay: sole, whiting, and 333 squids. These are coastal species with potentially shared spatio-temporal 334 dynamics e.g. same reproduction period (Alglave, Vermard, et al., 2023). 335 The temporally synthetic multivariate EOFs (EOFs on  $\mathbf{S}'^{(row)}_{multi}$ ) provides 336 seasonal loadings that represent the reproduction phenology for all the 337 species. We only retain the first two dimensions for the analysis (Figure 338 S4.1). For sole, the reproduction areas are generally consistent with those 339

identified when conducting univariate EOFs (Figure 4). For squids and
whitings, reproduction areas are observed along the Landes coast (45°N 1.5°W), on the Vendée coast (46.5°N - 2°W), and in the north of the Bay of
Biscay (47.5°N - 3°W).

The spatially synthetic multivariate EOFs is provided in SM S4.5 as well as 344 the lag EOFs (SM S4.6). The spatially synthetic representation allows to 345 identify the common spatial patterns among each three species as well as 346 synchrony in the spatio-temporal variation of the species. The lag EOFs is 347 applied to the NDVI case study. Compared with standard EOFs (Section 3), 348 it allows to identify an additional periodic signal with a 6-months phase 349 related to an alternance of high growth of vegetation in spring and autumn 350 and lower growth in summer and winter. 351

#### 352 6 Discussion

#### 353 Towards a wider use of EOFs in ecology

EOFs have tremendous potential for ecology. As massive and diverse datasets become available (satellite data, IST-SDM, satellite data and biogeochemical models), these decomposition methods will play a growing role in (1) reducing data dimensionality and (2) extracting ecologically relevant and interpretable information. In this paper, we demonstrate their potential for both terrestrial and marine systems. We argue that EOFs shouldbe systematically used to analyze spatio-temporal ecological datasets.

We presented two extensions of EOFs, one to better disentangle the processes identified through EOFs and improve ecological interpretability of the factor and another to apply EOFs to multivariate data. Many other methods derived from EOFs could be used to perform spatio-temporal analysis on ecological data and are available in the literature (Preisendorfer, 1988; Von Storch and Zwiers, 2002; Cressie and Wikle, 2011; Hannachi, 2021).

#### 368 Accounting for spatio-temporal correlation in EOFs

One might argue that EOFs do not explicitly account for spatial or temporal correlations and then are not per se a spatio-temporal method. It is true that EOFs is basically a PCA representation of spatio-temporal data; the spatio-temporal aspects of the method comes from the data rather than from the mathematical formulation of the method per se.

In a spatial context, methods have been developed to better handle spatial
correlations and to propose multiscale decompositions of ecological data
(P. Legendre and L. Legendre, 2012). Specifically, 'spatial eigenfunction
analysis' (e.g., Moran's eigenvector maps, asymmetric eigenvector maps,
multiscale ordination) estimate sets of eigenvectors based on spatial

configuration matrices. These methods could potentially be extended to
spatio-temporal data to identify distinct spatio-temporal scales that structure
ecological data.

More recently, Petitgas, Renard, et al. (2020) and Bez, Renard, and
Ahmed-Babou (2023) have proposed a method (Min-max Autocorrelation
Factor - MAF - or Empirical Orthogonal Maps - EOM) to provide spatially
decorrelated factors. EOM provide maps with stronger orthogonality
constraints and aims to better separate the information carried by the factors.
Still, such approach is very sensitive to the choice of the spatial lag chosen
in the analysis and does not account for temporal correlation.

# 389 Non-stationary, non-linear EOFs for spatio-temporal analysis under

### 390 climate change

Finally, some important hypotheses underlie EOFs, namely stationarity and 391 linearity (Hannachi, Jolliffe, and Stephenson, 2007; Alglave, Olmos, et al., 392 2024). EOFs can not be used to describe propagating patterns such as range 393 expansion or contraction as is common in species redistribution (Melles 394 et al., 2011; Scheele et al., 2017). Also, as in standard PCA, all non-linear 395 processes will not be detected and captured through EOFs while it is largely 396 recognized that ecosystems are strongly structured by non-linear 397 relationships and sometimes switching dynamics (Scheffer et al., 2001). 398 More generally speaking, handling non-stationarity in a changing 399

400 environment is a key challenge for statistical ecology (Litzow et al., 2019;

401 Astigarraga et al., 2020; Bueno de Mesquita et al., 2021).

Although some techniques, such as Hilbert complex EOFs or non-linear 402 PCA (Esquivel and Messina, 2008; Bueso, Piles, and Camps-Valls, 2020), 403 allow to handle non-linear relationship and to evidence propagating patterns, 404 these have been scarcely used in practice to investigate the effect of climate 405 change on ecosystems. Developing and applying approaches to handle 406 non-linearity and non-stationarity is critical in analysing the effect of 407 climate change on ecosystems; it constitues an open research avenue for 408 future study to track the effect of climate change on ecosystems. 409

#### 410 Data availability statement

411 Codes and data are available through the link here.

## 412 Acknowledgments

The authors are grateful to Julia Indivero (School of Aquatic and Fishery
Sciences, University of Washington) and Nikolai Morokhovich (School of
Aquatic and Fishery Sciences, University of Washington) who made the
internal review for the NOAA pre-submission procedure, as well as Nicolas

417 Bez who provided insightful comments on the manuscript. The findings and

418 conclusions of the present paper are those of the authors.

## 419 **Competing interests**

- 420 The authors declare that they have no conflicts of interest or competing
- 421 interests that could influence the content or interpretation of this manuscript.

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Fig. 1: NDVI data. (left) Monthly spatial data of NDVI S(x, t). (right) Monthly anomalies of the spatial predictions  $S^*(x, t)$ . Each panel corresponds to the average predictions or anomalies for a month over the period 2000 - 2023.



Fig. 2: EOFs results for satellite NDVI data (left) and IST-SDM predictions (right). (Top) Factors for the two first dimensions of the EOFs. (Bottom) Loadings for the two first dimensions of the EOFs. Blue dashed vertical lines correspond to the month of January for each year. The third and fourth dimensions of the IST-SDM case study are presented in Figure S4.2.



Fig. 3: Satellite NDVI data. Results of the canonical correlation analysis. (Top left) Spatial basis vector that maximize the correlation between the temporal variables  $y_1$  and  $y_2$ . (Top right) Correlation matrix between the two first EOFs loadings, the CCA  $y_1$ , and the ancillary variable. (Bottom) Comparison of the EOFs loadings with the ecological ancillary variable, and the CCA variable with the ecological ancillary variable. Blue dashed vertical lines correspond to the month of January for each year.



Fig. 4: IST-SDM predictions. Results for the multivariate temporally synthetic EOFs (EOFs on  $\mathbf{S}'_{multi}^{(row)}$ ). (Top) Factor maps for each species and first two dimensions. (Bottom) Loadings for the first two dimensions. Blue dashed vertical lines correspond to the month of January for each year. The loadings emphasize the seasonality of the reproduction for each three species. Orange areas in both dimensions correspond to the reproduction grounds.