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To cite this version:

Mengyuan Zhao, Maël Le Treust, Tobias Oechtering. Coordination Coding with Causal Encoder for Vector-Valued Witsenhausen Counterexample. ISIT 2024 - IEEE International Symposium on Information Theory, Jul 2024, Athens, Greece. pp.3255-3260, 10.1109/ISIT57864.2024.10619377. hal-04679217

HAL Id: hal-04679217 <https://hal.science/hal-04679217v1>

Submitted on 27 Aug 2024

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Coordination Coding with Causal Encoder for Vector-valued Witsenhausen Counterexample

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Abstract—We investigate the Witsenhausen counterexample in a continuous vector-valued context with a causal encoder and noncausal decoder. Our main result is the optimal single-letter condition that characterizes the set of achievable Witsenhausen power costs and estimation costs, leveraging a modified weak typicality approach. In particular, we accommodate our power analysis to the causal encoder constraint, and provide an improved distortion error analysis for the challenging estimation of the interim state. Interestingly, the idea of dual role of control is explicitly captured by the two auxiliary random variables.

I. INTRODUCTION

In 1968, Wistenhausen proposed a famous counterexample [1] illustrating the challenges posed by non-classical information pattern, for which affine functions are strictly suboptimal in LQG setting. Nowadays, it still plays an important role in contemporary distributed decision making [2]–[4], and information-theoretic control fields [5]–[11].

Recent advancements in studying the Witsenhausen counterexamples have been driven by information theory and coding considering its vector-valued extensions. This enables the adoption of advanced coding techniques such as [12]– [14]. Pioneering this approach, Grover et al. explored a setting in which both decision-makers are assumed to be non-causal [15].

Empirical coordination, [16]–[19], to mention a few, has emerged as a powerful tool in characterizing the cost and capacity trade-off for state-dependent channels when the cost function jointly evaluates the source symbol and the channel input symbol. This method has shown its effectiveness in causal and non-causal encoding and decoding frameworks, such as [20, 21], and even in distributed decision making problem with power control [22]. Utilizing these insights, [23] provides the optimal characterizations for finite alphabet vector Witsenhausen problems with causal and non-causal decisionmakers. Moreover, an achievability proof based on empirical coordination coding [24] for the non-causal encoding and causal decoding setup processing Gaussian states [25] is presented in [26, 27].

Diverging from the setting in [27], this paper explores the scenario where the first decision-maker encodes the state sequence causally, while the subsequent one estimates the interim state non-causally. Since some random variables are continuous, a modified version of weak typicality [28] is required in order to obtain a desired Markov Lemma-like result, as done in [29]. Moreover, to implement the Witsenhausen

Fig. 1. The i.i.d. state and the channel noise are drawn according to Gaussian distributions $X_0^n \sim \mathcal{N}(0, Q\mathbb{I})$ and $Z_1^n \sim \mathcal{N}(0, N\mathbb{I})$.

cost constraints, we impose an average power constraint and provide a simplified approach to analyse the distortion cost. The converse proof validates the optimality of our set of the achievable pair of costs.

II. SYSTEM MODEL

Our vector-valued Witsenhausen setup consists of *causal* source states and channel noises that are drawn independently according to the i.i.d. Gaussian distributions $X_0^n \sim \mathcal{N}(0, Q\mathbb{I})$ and $Z_1^n \sim \mathcal{N}(0, N\mathbb{I})$, for some $Q, N > 0$. We denote by X_1 the memoryless interim state and Y_1 and output of the memoryless additive channel, generated by

$$
X_1 = X_0 + U_1 \qquad \text{with } X_0 \sim \mathcal{N}(0, Q), \text{ (1)}
$$

$$
Y_1 = X_1 + Z_1 = X_0 + U_1 + Z_1 \quad \text{with } Z_1 \sim \mathcal{N}(0, N). \text{ (2)}
$$

We denote by $\mathcal{P}_{X_0} = \mathcal{N}(0, Q)$ the generative Gaussian probability distribution of the state, and by $\mathcal{P}_{X_1Y_1|X_0U_1}$ the channel probability distribution according to (1) and (2). Work [27] considered noncausal encoder paired with causal decoder. In this work, we consider a setup with flipped causality:

Definition II.1. *For* $n \in \mathbb{N}$, *a* "control design" with causal *encoder and noncausal decoder is a tuple of stochastic functions* $c = (\{f_{U_{1,t}|X_0^t}^{(t)}\}_{t=1}^n, g_{U_2^n|Y_1^n})$ *defined by*

$$
f_{U_{1,t}|X_0^t}^{(t)}: \mathcal{X}_0^t \longrightarrow \mathcal{U}_1, \quad g_{U_2^n|Y_1^n}: \mathcal{Y}_1^n \longrightarrow \mathcal{U}_2^n, \tag{3}
$$

which induces a distribution over sequences of symbols:

$$
\prod_{t=1}^{n} \mathcal{P}_{X_{0,t}} \prod_{t=1}^{n} f_{U_{1,t}|X_0^t}^{(t)} \prod_{t=1}^{n} \mathcal{P}_{X_{1,t},Y_{1,t}|X_{0,t},U_{1,t}} g_{U_2^n|Y_1^n}.
$$
 (4)

We denote by $C_e(n)$ the set of control designs with causal *encoder and non-causal decoder.*

Definition II.2. *We define the two long-run cost functions* $c_P(u_1^n) = \frac{1}{n} \sum_{t=1}^n (u_{1,t})^2$ and $c_S(x_1^n, u_2^n) = \frac{1}{n} \sum_{t=1}^n (x_{1,t} -$

 $(u_{2,t})^2$. The pair of costs $(P, S) \in \mathbb{R}^2$ is achievable if for all $\varepsilon > 0$ *, there exists* $\bar{n} \in \mathbb{N}$ *such that for all* $n \geq \bar{n}$ *, there exists a control design* $c \in \mathcal{C}_e(n)$ *such that*

$$
\mathbb{E}\Big[\big| P - c_P(U_1^n) \big| + \big| S - c_S(X_1^n, U_2^n) \big| \Big] \le \varepsilon. \tag{5}
$$

Remark 1. *Equation* (5) *is a stronger requirement than*

$$
|P - \mathbb{E}\left[c_P(U_1^n)\right]| + |S - \mathbb{E}\left[c_S(X_1^n, U_2^n)\right]| \le \varepsilon,\qquad(6)
$$

which was considered in [27]. Here, both notions lead to the same expression for the single-letter solution.

Theorem II.3 (Main result). *The pair of Witsenhausen costs* (P, S) *is achievable if and only if there exists a joint distribution over the random variables* $(X_0, W_1, W_2, U_1, X_1, Y_1, U_2)$ *that decomposes according to*

$$
\mathcal{P}_{X_0}\mathcal{P}_{W_1}\mathcal{P}_{W_2|X_0,W_1}\mathcal{P}_{U_1|X_0,W_1}\mathcal{P}_{X_1,Y_1|X_0,U_1}\mathcal{P}_{U_2|W_1,W_2,Y_1},\tag{7}
$$

such that

$$
I(W_1, W_2; Y_1) - I(W_2; X_0 | W_1) \ge 0,
$$
\n(8)

$$
P = \mathbb{E}[U_1^2], \qquad S = \mathbb{E}[(X_1 - U_2)^2],
$$
 (9)

where \mathcal{P}_{X_0} and $\mathcal{P}_{X_1,Y_1|X_0,U_1}$ are given Gaussian distributions, and W_1, W_2 are two auxiliary random variables.

Remark 2. *The probability distribution* (7) *satisfies:*

$$
\begin{cases}\nX_0 \text{ is independent of } W_1, \\
U_1 \to (X_0, W_1) \to W_2, \\
(X_1, Y_1) \to (X_0, U_1) \to (W_1, W_2), \\
U_2 \to (W_1, W_2, Y_1) \to (X_0, U_1, X_1).\n\end{cases}
$$
\n(10)

The first two properties are derived as consequences of causal encoding, where the achievability proof requires the Shannon strategy [30] for generating new codewords to the channel. The third Markov chain is due to the memoryless property of the channel and the processing order. The last Markov chain comes from the non-causal decoding and the symbolwise reconstruction.

Remark 3. *The information constraint* (8) *reformulates as*

$$
I(W_1, W_2; X_0) \le I(W_1, W_2; Y_1).
$$

The LHS corresponds to the lossy source coding involved in the quantization of X_0 *using* (W_1, W_2) *. The RHS stands for the channel coding rate of outputting* Y_1 *given the input pair* (W_1, W_2) *. In particular,* W_1 *assists in enabling reliable communication, while* W_2 *is used to coordinate with the state.*

The achievability proof of Theorem II.3 is presented in Section III and the converse proof is in Section IV.

III. ACHIEVABILITY PROOF OF THE MAIN RESULT

Preliminaries: We define the entropy and the weakly typical set for random vectors that combine discrete and continuous components as in [27]. We consider an arbitrary but fixed ε > 0 and assume the sequence $(X_0^n, W_1^n, W_2^n, U_1^n, X_1^n, Y_1^n, U_2^n)$ is generated i.i.d. according to a distribution that decomposes

as (7), with $P = \mathbb{E}[U_1^2], S = \mathbb{E}[(X_1 - U_2)^2]$. Let $\psi^{(n)}$: $\mathcal{X}_0^n \times \mathcal{W}_1^n \times \mathcal{W}_2^n \times \mathcal{U}_1^n \times \mathcal{X}_1^n \times \mathcal{Y}_1^n \times \mathcal{U}_2^n \to \{0, 1\}$ denote an indicator function for sequences of length n with

$$
\psi^{(n)}(x_0^n, w_1^n, w_2^n, u_1^n, x_1^n, y_1^n, u_2^n)
$$
\n
$$
= \begin{cases}\n1 & \text{if } |c_S(x_1^n, u_2^n) - S| \ge \frac{1}{12}\varepsilon \\
& \text{or } (w_1^n, w_2^n, y_1^n) \notin \mathcal{A}_{\varepsilon}^n(W_1, W_2, Y_1), \\
0 & \text{otherwise.} \n\end{cases}
$$

Using the weak law of large numbers (LLN) and the union bound we have

$$
\mathbb{E}[\psi^{(n)}(X_0^n, W_1^n, W_2^n, U_1^n, X_1^n, Y_1^n, U_2^n)] \le \delta_n \to 0,
$$

as $n \to \infty$. And similarly, we define the set

$$
\mathcal{S}_{\varepsilon}^{(n)} = \{ (x_0^n, w_1^n, w_2^n) \mid \eta^{(n)}(x_0^n, w_1^n, w_2^n) \leq \sqrt{\delta_n} \},\
$$

where $\eta^{(n)}(x_0^n, w_1^n, w_2^n) = \mathbb{E}[\psi^{(n)}(x_0^n, w_1^n, w_2^n, U_1^n, X_1^n, Y_1^n,$ U_2^n | $X_0^n = x_0^n$, $W_1^n = w_1^n$, $W_2^n = w_2^n$]. We take $\mathcal{B}_{\varepsilon}^{(n)}$ = $\mathcal{A}_{\varepsilon}^{(n)} \cap \mathcal{S}_{\varepsilon}^{(n)}$. Then, we can easily prove the subsequent lemma:

Lemma III.1. Let the sequence (X_0^n, W_1^n, W_2^n) i.i.d. \sim $\mathcal{P}_{X_0,W_1,W_2}$, then

$$
\mathbb{P}\left(\left(X_0^n, W_1^n, W_2^n\right) \in \mathcal{B}_{\varepsilon}^{(n)}\right) \xrightarrow{n \to \infty} 1.
$$

Proof of Lemma III.1. From the Markov inequality we obtain

$$
\mathbb{P}\left((X_0^n, W_1^n, W_2^n) \notin \mathcal{S}_{\varepsilon}^{(n)}\right) \leq \frac{\delta_n}{\sqrt{\delta_n}} = \sqrt{\delta_n}.
$$

Lemma III.1 elaborates that the certainty of an outcome being jointly weakly typical AND satisfying the distortion cost constraint is guaranteed, as per the AEP, if it is generated i.i.d.

Moreover, it is also necessary to quantize the output of the non-causal decoder, as done in [27], to ensure that the second cost constraint remains bounded in case an error happens. Given a joint distribution with $\mathbb{E}[(X_1 - U_2)^2] = S$, for any $\hat{\delta} > 0$ there exists a quantization $q_{U_2} : \mathcal{U}_2 \to {\hat{u}_{2,k}}_{k=1}^{N_{U_2}}$, as explained in [31], such that

$$
\hat{S} = \mathbb{E}[(X_1 - q_{U_2}(U_2))^2] \le (1 + \hat{\delta})S,
$$

in particular $\hat{\delta}S < \frac{1}{4}\varepsilon$.

Now, we provide a coding scheme by considering a probability distribution $\mathcal{P}_{X_0,W_1,W_2,U_1,X_1,Y_1,U_2}$ that decomposes as (7) such that there exists a $\delta > 0$, and a rate $R > 0$, satisfying $I(W_2; X_0|W_1) + \delta \le R \le I(W_1; Y_1) + I(W_2; Y_1|W_1) - \delta.$ We consider a block-Markov code $c \in \mathcal{C}_e(nB)$ defined over $B \in \mathbb{N}$ blocks of length *n*, similar to the scheme in [14].

Random codebook generation: We generate $|M| = 2^{nR}$ sequences $W_1^n(m)$ i.i.d. $\sim \mathcal{P}_{W_1}$ with index $m \in \mathcal{M}$. For each index $m \in \mathcal{M}$, we generate the same number $|\mathcal{M}| = 2^{nR}$ sequences $W_2^n(m, \hat{m})$ with index $\hat{m} \in \mathcal{M}$ i.i.d. ~ $\mathcal{P}_{W_2|W_1}$ depending on sequence $W_1^n(m)$.

Encoding function: Due to the causal feature, the encoder performs *forward encoding*, where the encoder outputs the information associated with the *past* block. Let m_b denote the message generated during block $b \in [1 : B]$. During the first block, without loss of generality, the encoder takes $m_1 = 1$ and returns $W_1^n(m_1)$. At the beginning of block $b \in [2 : B]$, the encoder has observed the full sequence of source symbols $X_{0,b-1}^n \in \mathcal{X}^n$ of the previous block. It recalls the index $m_{b-1} \in \mathcal{M}$ of the sequence $W_1^n(m_{b-1})$ used for block $b-1$. It finds an index m_b such that sequences $(X_{0,b-1}^n, W_1^n(m_{b-1}), W_2^n(m_{b-1}, m_b)) \in \mathcal{B}_{\varepsilon}^{(n)}(\mathcal{P}_{X_0,W_1,W_2})$ are jointly typical. It returns $W_1^n(m_b)$ corresponding to the current block b, and for each time $t \in [1:n]$, it sends the symbol $U_{1,t,b}$ i.i.d. $\sim \mathcal{P}_{U_1|X_0,W_1}$ depending on $W_{1,t}(m_b)$ and $X_{0,t,b}$ causally observed in the current block $b \in [1 : B]$.

Decoding function: The decoder first returns $\tilde{m}_1 = 1$. During block $b \in [2 : B]$, the decoder recalls the past sequence $Y_{1,b-1}^n$ and the index \tilde{m}_{b-1} that corresponds to the sequence $\tilde{W}_{1,b-1}^n = W_1^n(\tilde{m}_{b-1})$. It observes the channel output $Y_{1,b}^n$ and finds the unique index \tilde{m}_b such that $(Y_{1,b}^n, W_1^n(\tilde{m}_b)) \in$ $\mathcal{A}_\varepsilon^{(n)}(\mathcal{P}_{Y_1,W_1})$ and $(Y_{1,b-1}^n,W_1^n(\tilde{m}_{b-1}),W_2^n(\tilde{m}_{b-1},\tilde{m}_{b})) \in$ $\mathcal{A}_{\varepsilon}^{(n)}(\mathcal{P}_{Y_1,W_1,W_2})$. We denote by $\tilde{W}_{1,b}^n = W_1^n(\tilde{m}_b)$ and $\tilde{W}_{2,b-1}^n = W_2^n(\tilde{m}_{b-1}, \tilde{m}_b)$ as our choice.

Forward transmission of the decoder: Upon receiving all sequences for all first $(B-1)$ blocks, the decoder non-causally generates $U_{2,b}^n \sim \mathcal{P}_{U_2|Y_1,W_1,W_2}^{\otimes n}$ depending on sequences $(Y_{1,b}^n, \tilde{W}_{1,b}^n, \tilde{W}_{2,b}^n)$ for $b \in [1 : B - 1]$. Finally, the decoder transmits the quantized sequence $\hat{U}_{2,b}^n$ as the reconstruction for the interim state $X_{1,b}^n$.

Termination block B*:* As for the last block, the decoder simply outputs an all zero sequence, i.e., $\hat{U}_{2,B}^n = 0$. Usually, sequences are *not* jointly typical in the last block. Therefore we omit its error analysis in the subsequent part.

Error analysis per block: We first focus on the encoding error \mathcal{E}^e . For $b \in [2 : B]$, let $\mathcal{E}_b^e(m_{b-1})$ denote the event of a failed encoding process during block *b* given the knowledge of m_{b-1} , i.e., $\mathcal{E}_b^e(m_{b-1})$ = $\{\forall m_b \in \mathcal{M}, (X_{0,b-1}^n, W_1^n(m_{b-1}), W_2^n(m_{b-1}, m_b)) \notin$ $\mathcal{B}_{\varepsilon}^{(n)}(\mathcal{P}_{X_0,W_1,W_2})\}$. Due to the independence between the codewords and Markov blocks, the probability of an encoding error in block b given no encoding errors in the previous blocks is

$$
\mathbb{P}(\mathcal{E}_b^e(M_{b-1}) \mid \bigcap_{\beta=2}^{b-1} \bar{\mathcal{E}}_\beta^e(M_{\beta-1})) = \mathbb{P}(\mathcal{E}_b^e(M_{b-1}))
$$

If $R \ge I(W_2; X_0|W_1) + \delta$, following the Covering Lemma [32] and Lemma III.1, we have $\mathbb{P}\left(\mathcal{E}_{b}^{e}(M_{b-1})\right) \to 0$ as $n \to \infty$, $\forall b \in [2 : B]$. Thus, state sequences for the first $B - 1$ blocks are successfully encoded with probability $\rightarrow 1$ as $n \rightarrow \infty$.

Next, we analyze the decoding error \mathcal{E}^d . For $b \in [2 : 1]$ B], let $\mathcal{E}_{b}^{d}(m_{b-1})$ denote the event of a failed decoding process in block b given the past estimated index m_{b-1} . Furthermore, let $\mathcal{E}_{b-1}^{d,\bar{t},1}(m_{b-1})$ denote the event that the sequence $Y_{1,b-1}^n$ given m_{b-1} is not jointly typical, i.e., $\mathcal{E}_{b-1}^{d,t,1}(m_{b-1}) = \{ (Y_{1,b-1}^n, W_{1,b-1}^n(m_{b-1}), W_2^n(m_{b-1}, M_b)) \notin$ $\mathcal{A}_{\varepsilon}^{(n)}(\mathcal{P}_{Y_1,W_1,W_2})\}$ and let $\mathcal{E}_{b}^{d,t,2}$ denote that the current sequence $Y_{1,b}^n$ is not jointly typical, i.e., $\mathcal{E}_b^{d,t,2}$ = $\{(Y_{1,b}^n, W_1^n(M_b)) \notin \mathcal{A}_{\varepsilon}^{(n)}(\mathcal{P}_{Y_1,W_1})\}$. Note that, if a sequence with fewer terms $(Y_{1,b}^n, W_1^n(m_b))$ is atypical, it implies that a sequence with more terms $(Y_{1,b}^n, W_1^n(m_b), W_2^n(m_b, m_{b+1}))$ is also atypical. Therefore, we have $\mathcal{E}_b^{d,t,2} \subset \mathcal{E}_b^{d,t,1}(M_b)$. Then, the decoding error probability given no past decoding error or encoding error $\mathbb{P}(\mathcal{E}_b^d(m_{b-1}) \mid \bigcap_{\beta=2}^{b-1} \mathcal{E}_\beta^d(m_{\beta-1}) \cap \overline{\mathcal{E}}^e)$ can be upperbounded by

$$
\mathbb{P}(\mathcal{E}_b^d(m_{b-1}) \mid \cap_{\beta=2}^{b-1} \bar{\mathcal{E}}_\beta^d(m_{\beta-1}) \cap \bar{\mathcal{E}}^e \cap \bar{\mathcal{E}}_{b-1}^{d,t,1}(m_{b-1}) \cap \bar{\mathcal{E}}_b^{d,t,2}) + \mathbb{P}(\mathcal{E}_{b-1}^{d,t,1}(m_{b-1}) \cup \mathcal{E}_b^{d,t,2} \mid \cap_{\beta=2}^{b-1} \bar{\mathcal{E}}_\beta^d(m_{\beta-1}) \cap \bar{\mathcal{E}}^e) \tag{11}
$$

using the union bound. The first term of (11) can be upperbounded by

$$
\mathbb{P}(\mathcal{E}_b^d(m_{b-1}) \mid \bigcap_{\beta=2}^{b-1} \bar{\mathcal{E}}_{\beta}^d(m_{\beta-1}) \cap \bar{\mathcal{E}}^e \cap \bar{\mathcal{E}}_{b-1}^{d,t,1}(m_{b-1}) \cap \bar{\mathcal{E}}_b^{d,t,2})
$$
\n
$$
= \mathbb{P}\left(\exists m' \neq M_b, \text{ s.t. } \{W_1^n(m') \in \mathcal{A}_{\varepsilon}^{(n)}(W_1|Y_{1,b}^n)\} \cap \{W_2^n(M_{b-1}, m') \in \mathcal{A}_{\varepsilon}^{(n)}(W_2^n|Y_{1,b-1}^n, W_{1,b-1}^n)\right)
$$
\n(a)

where (a) is due to the joint Packing Lemma [32] in order to satisfy both conditions at the same time given that $R <$ $I(W_1; Y_1) + I(W_2; Y_1|W_1) - 7\varepsilon$. Moreover, as for the second term, we have

$$
\mathbb{P}(\mathcal{E}_{b-1}^{d,t,1}(m_{b-1}) \cup \mathcal{E}_{b}^{d,t,2} | \cap_{\beta=2}^{b-1} \bar{\mathcal{E}}_{\beta}^{d}(m_{\beta-1}) \cap \bar{\mathcal{E}}^{e})
$$
\n
$$
\stackrel{\text{(b)}}{\leq} \mathbb{P}((Y_{1,b-1}^{n}, W_{1,b-1}^{n}, W_{2,b-1}^{n}) \notin \mathcal{A}_{\varepsilon}^{(n)} |
$$
\n
$$
(X_{0,b-1}^{n}, W_{1,b-1}^{n}, W_{2,b-1}^{n}) \in \mathcal{B}_{\varepsilon}^{(n)})
$$
\n
$$
+ \mathbb{P}((Y_{1,b}^{n}, W_{1,b}^{n}, W_{2,b}^{n}) \notin \mathcal{A}_{\varepsilon}^{(n)} | (X_{0,b}^{n}, W_{1,b}^{n}, W_{2,b}^{n}) \in \mathcal{B}_{\varepsilon}^{(n)})
$$
\n
$$
\stackrel{\text{(c)}}{\leq} 2 \cdot \max_{(x_{0}^{n}, w_{1}^{n}, w_{2}^{n}) \in \mathcal{B}_{\varepsilon}^{(n)}} \eta^{(n)}(x_{0}^{n}, w_{1}^{n}, w_{2}^{n})
$$
\n
$$
\leq 2\sqrt{\delta_{n}} \xrightarrow{n \to \infty} 0,
$$

where (b) comes from the independence of each Markov block and the fact that $\{(Y_{1,b}^n, W_{1,b}^n) \notin \mathcal{A}_{\varepsilon}^{(n)}\} \subset$ $\{(Y_{1,b}^n, W_{1,b}^n, W_{2,b}^n) \notin \mathcal{A}_{\varepsilon}^{(n)}\}\$, and (c) comes from the definition of $\mathcal{B}_{\varepsilon}^{(n)}$.

Therefore, following the above arguments, the encoding error and the decoding error are both asymptotically zero.

Witsenhausen costs analysis: We first analyze the power cost. To implement the power constraint, we impose a meansquare transmission power constraint as done in [14, 15]:

$$
\sum_{t=1}^{n} \mathbb{E}\left[U_{1,t}^{2}\right] \leq nP,\tag{12}
$$

where the expectation is taken over random sequences X_0^n and W_1^n . Therefore, the generated sequence U_1^n satisfies the condition of uniform integrability. Thus, for $b \in [1 : B]$, since at each time $U_{1,t,b}$ is generated i.i.d. with finite second moment $\mathbb{E}\left[U_1^2\right] = P$, due to the LLN, the non-negative averaged power sequence $c_P(U_{1,b}^n) = \frac{1}{n} \sum_{t=1}^n (U_{1,t,b}^2)$ converges in probability to P. Moreover, since sequence $U_{1,b}^{n}$ is also uniformly integrable, it holds (stronger) that

$$
\mathbb{E}\left[|c_P(U_{1,b}^n) - P|\right] < \varepsilon, \quad b \in [1:B] \tag{13}
$$

for *n* sufficiently large, more details in [33].

Next, we examine the estimation cost. Define the indicator function $\rho_b = 0$ to denote the absence of coding error in block $b \in [1 : B - 1]$. Now, take $\phi_b = (1 - \psi_b^{(n)})$ $b^{(n)}(1-\rho_b)$ to indicate the occurrence of desired sequences that satisfy both the estimation cost and the typicality constraints AND no coding error event in block *b*. From the LLN and the error analysis, we have $\mathbb{E} [\phi_b] \to 1$ as $n \to \infty$ for $b \in [1 : B - 1]$. In particular, if $\phi_b = 1$, we have $\mathbb{E}\left[|c_S(X_{1,b}^n, \hat{U}_{2,b}^n) - S| \mid \phi_b = 1\right] < \frac{1}{12}\varepsilon$. Therefore, from the law of total expectation, we have

$$
\mathbb{E}\left[|c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n}) - \hat{S}|\right]
$$
\n
$$
= \mathbb{P}\left(\phi_{b} = 1\right) \cdot \mathbb{E}\left[|c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n}) - \hat{S}|\ \big|\ \phi_{b} = 1\right]
$$
\n
$$
+ \mathbb{P}\left(\phi_{b} = 0\right) \cdot \mathbb{E}\left[|c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n}) - \hat{S}|\ \big|\ \phi_{b} = 0\right]
$$
\n
$$
\leq \frac{1}{12}\varepsilon + \mathbb{P}\left(\phi_{b} = 0\right) \cdot \hat{S}
$$
\n
$$
+ \mathbb{P}\left(\phi_{b} = 0\right) \cdot \mathbb{E}\left[c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n})\ \big|\ \phi_{b} = 0\right].
$$

For *n* sufficiently large, the second term $\mathbb{P}(\phi_b = 0) \cdot \hat{S} \leq$ $\frac{1}{12}\varepsilon$, since $\hat{S} < \infty$. Next, we evaluate the last term above. In particular, we need $c_S(X_{1,b}^n, \hat{U}_{2,b}^n)$ to be stochastically bounded when an unwanted error event happens. Since

$$
\mathbb{P}\left(\phi_b=0\right) \cdot \mathbb{E}\left[c_S(X_{1,b}^n, \hat{U}_{2,b}^n) \middle| \phi_b=0\right]
$$
\n(14)

$$
= \mathbb{P}\left(\phi_b = 0\right) \cdot \mathbb{E}\left[\frac{1}{n}\sum_{t=1}^n |X_{0,t,b} + U_{1,t,b} - \hat{U}_{2,t,b}|^2 \middle| \phi_b = 0\right],
$$

by applying the inequality $(a + b)^2 \le 2a^2 + 2b^2$, $\forall a, b \in \mathbb{R}$, we have

$$
\mathbb{P}(\phi_b = 0) \cdot \mathbb{E} \left[c_S(X_{1,b}^n, \hat{U}_{2,b}^n) \middle| \phi_b = 0 \right]
$$

\n
$$
\leq 2\mathbb{P}(\phi_b = 0) \cdot \mathbb{E} \left[c_P(U_{1,b}^n) \middle| \phi_b = 0 \right]
$$

\n
$$
+ 2\mathbb{P}(\phi_b = 0) \cdot \mathbb{E} \left[c_S(X_{0,b}^n, \hat{U}_{2,b}^n) \middle| \phi_b = 0 \right]
$$

\n
$$
\leq 2\mathbb{P}(\phi_b = 0) \cdot \mathbb{E} \left[c_P(U_{1,b}^n) - P \middle| + P \middle| \phi_b = 0 \right]
$$

\n
$$
+ 2\mathbb{P}(\phi_b = 0) \cdot \mathbb{E} \left[c_S(X_{0,b}^n, \hat{U}_{2,b}^n) \middle| \phi_b = 0 \right]
$$

\n
$$
= 2P \mathbb{P}(\phi_b = 0) + 2\mathbb{P}(\phi_b = 0) \cdot \mathbb{E} \left[c_P(U_{1,b}^n) - P \middle| \phi_b = 0 \right]
$$

\n
$$
+ 2\mathbb{P}(\phi_b = 0) \cdot \mathbb{E} \left[c_S(X_{0,b}^n, \hat{U}_{2,b}^n) \middle| \phi_b = 0 \right].
$$
 (15)

Here, the first term $2P\mathbb{P}(\phi_b = 0) \leq \frac{1}{36}\varepsilon$ since $P < \infty$. And the second term $2\mathbb{P}\left(\phi_b=0\right) \cdot \mathbb{E}\left[\left|c_P(U_{1,b}^n)-P\right|\Big|\phi_b=0\right] \leq$ $2\mathbb{E}\left[|c_P(U_{1,b}^n)-P|\right] \leq \frac{1}{36}\varepsilon$ due to the nonnegativity of $|c_P(U_{1,b}^n) - P|$ and the power constraint (13).

Next, we analyse the last term in (15) with Wyner's trick [31] exploiting the discretization of U_2 as follows

$$
\mathbb{E}\left[c_S(X_{0,b}^n, \hat{U}_{2,b}^n) \middle| \phi_b = 0\right]
$$

$$
\leq \frac{1}{n} \sum_{t=1}^n \mathbb{E}\left[D(X_{0,t,b}) \middle| \phi_b = 0\right],
$$

with $D(X_{0,t,b}) = \max_{\hat{u}_{2,k}} c_S(X_{0,t,b}, \hat{u}_{2,k})$. The random variables $\{D(X_{0,t,b})\}_{t=1}^n$ are i.i.d. and integrable since $c_S(\cdot, \cdot)$ is a quadratic measure and $X_{0,t,b}$ is i.i.d. Gaussian distributed. Next, let $\chi_{\{D(X_{0,t,b})>d\}}$ denote an indicator function indicating if $D(X_{0,t,b}) > d$ for some $0 < d < \infty$. Then we have

$$
\mathbb{P}(\phi_b = 0) \cdot \mathbb{E}\left[c_S(X_{0,b}^n, \hat{U}_{2,b}^n) \middle| \phi_b = 0\right]
$$
\n
$$
\leq d \cdot \mathbb{P}(\phi_b = 0) + \mathbb{E}\left[D(X_{0,t,b})\chi_{\{D(X_{0,t,b}) > d\}}\right].
$$
\n(16)

Given that $D(X_{0,t,b})$ is integrable, for any $\varepsilon_d > 0$, there must exist a d_0 such that $\mathbb{E}\left[D(X_{0,t,b})\chi_{\{D(X_{0,t,b})>d\}}\right]<\varepsilon_d$ for all $d > d_0$, as per the Monotone Convergence Theorem. Thus, for a sufficiently small ε_d and a sufficiently large n, term (16) can be upper bounded by $\frac{1}{72}\varepsilon$, so that $(14) \le \frac{1}{12}\varepsilon$.

Thus, compared to the original cost constraint S, for block $b \in [1 : B - 1]$ we have

$$
\mathbb{E}\left[|c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n}) - S|\right] \leq |\hat{S} - S| + \mathbb{E}\left[|c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n}) - \hat{S}|\right] \leq \frac{1}{4}\varepsilon + \mathbb{E}\left[|c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n}) - \hat{S}|\right] \leq \frac{1}{2}\varepsilon.
$$
 (17)

Lastly, we estimate the cost of the last block where the output $\hat{U}_{2,B}^n$ is set to be an all-zero sequence. In this case

$$
\mathbb{E}\left[|c_{S}(X_{1,B}^{n}, \hat{U}_{2,B}^{n}) - \hat{S}|\right] \n\leq \mathbb{E}\left[|2c_{P}(X_{0,B}^{n}) + 2c_{P}(U_{1,B}^{n}) - \hat{S}|\right] \n\leq 2 \cdot \mathbb{E}\left[c_{P}(X_{0,B}^{n})\right] + 2 \cdot \mathbb{E}\left[c_{P}(U_{1,B}^{n})\right] + \hat{S} \n\stackrel{\text{(a)}}{\leq} 2 \cdot Q + 2 \cdot P + \hat{S} < \infty,
$$

where (a) comes from that $X_{0,B}^n$ is generated i.i.d. ~ $\mathcal{N}(0, Q\mathbb{I})$. Therefore, similar to (17), we have

$$
\mathbb{E}\left[|c_{S}(X_{1,B}^{n}, \hat{U}_{2,B}^{n}) - S|\right] \leq \frac{1}{4}\varepsilon + 2 \cdot Q + 2 \cdot P + \hat{S} \triangleq \tilde{S}.
$$

Synthesizing power costs for all B blocks, we have

$$
\mathbb{E}\left[|c_P(U_1^{Bn}) - P|\right] \le \frac{1}{B} \sum_{b=1}^B \mathbb{E}\left[|c_P(U_{1,b}^n) - P|\right] \le \varepsilon, \tag{18}
$$

for n sufficiently large.

For the average estimation cost, the impact of the last pathological block can be effectively diminished, since

$$
\mathbb{E}\left[|c_{S}(X_{1}^{Bn}, \hat{U}_{2}^{Bn}) - S|\right] \leq \frac{1}{B} \sum_{b=1}^{B} \mathbb{E}\left[|c_{S}(X_{1,b}^{n}, \hat{U}_{2,b}^{n}) - S|\right] \leq \frac{1}{B} \tilde{S} + \frac{B-1}{B} \cdot \frac{1}{2} \varepsilon \leq \varepsilon,
$$
\n(19)

for n and B sufficiently large.

Existence of the desired code: With the aid of random coding arguments and applying the Selection Lemma [34] to both (18) (19), we can obtain that there exists a control design $c \in \mathcal{C}_e(n)$ that simultaneously satisfies

$$
\mathbb{E}\left[|c_P(U_1^{Bn}) - P|\right] < \varepsilon, \quad \mathbb{E}[|c_S(X_1^{Bn}, \hat{U}_2^{Bn}) - S|] < \varepsilon.
$$

Closedness: So far we have only shown under strict inequality of rate constraint (8) the cost constraints holds. As for equality, we consider a sequence of probability distribution $\mathcal{P}^k \xrightarrow{k \to \infty} \mathcal{P}^*$, where each \mathcal{P}^k satisfies the rate constraint with strict inequality, and \mathcal{P}^* satisfies it with equality. Therefore,

$$
\mathbb{E}_{\mathcal{P}^*}[|c_S(X_1^n, U_2^n) - S|]
$$

\n
$$
\leq \mathbb{E}_{\mathcal{P}_k}[|c_S(X_1^n, U_2^n) - S|] + \mathbb{E}_{\mathcal{P}^* - \mathcal{P}_k}[|c_S(X_1^n, U_2^n) - S|]
$$

\n
$$
\leq 1/2 \cdot \varepsilon + 1/2 \cdot \varepsilon \leq \varepsilon
$$

where the first $1/2 \cdot \varepsilon$ comes from the achievability proof with strict inequality as n large enough, and the second $1/2 \in \text{could}$ similarly be derived repeating the estimation cost analysis steps above since the probability measure $\rightarrow 0$ as $k \rightarrow 0$. Moreover, the power constraint for \mathcal{P}^* also holds due to (12).

IV. CONVERSE PROOF OF THE MAIN RESULT

We consider a control design $c \in \mathcal{C}_e(n)$ of block length $n \in$ $\mathbb N$ that achieves a pair of costs (P, S) . According to Csiszár sum identity, we have

$$
0 = \sum_{t=1}^{n} I\left(X_{0}^{t-1}; Y_{1,t}|Y_{1,t+1}^{n}\right) - \sum_{t=1}^{n} I\left(Y_{1,t+1}^{n}; X_{0,t}|X_{0}^{t-1}\right)
$$

\n
$$
\leq \sum_{t=1}^{n} I\left(X_{0}^{t-1}, Y_{1,t+1}^{n}; Y_{1,t}\right) - \sum_{t=1}^{n} I\left(Y_{1,t+1}^{n}; X_{0,t}|X_{0}^{t-1}\right)
$$

\n
$$
\stackrel{(a)}{=} \sum_{t=1}^{n} I(W_{1,t}, W_{2,t}; Y_{1,t}) - \sum_{t=1}^{n} I(W_{2,t}; X_{0,t}|W_{1,t})
$$

\n
$$
\stackrel{(b)}{=} n \cdot (I(W_{1,T}, W_{2,T}; Y_{1,T}|T) - I(W_{2,T}; X_{0,T}|W_{1,T}, T))
$$

\n
$$
\stackrel{(c)}{=} n \cdot (I(W_{1}, T, W_{2,T}; Y_{1,T}) - I(W_{2,T}; X_{0,T}|W_{1,T}, T))
$$

\n
$$
\stackrel{(c)}{=} n \cdot (I(W_{1}, W_{2}; Y_{1}) - I(W_{2}; X_{0}|W_{1})),
$$

where (a) is from the identification of the auxiliary random variables $W_{1,t} = X_0^{t-1}$, $W_{2,t} = Y_{1,t+1}^n$, (b) is from introducing the uniform random variable $T \in \{1, ..., n\}$ and the corresponding mean auxiliary random variables $X_{0,T}$, $W_{1,T}$, $W_{2,T}$, $Y_{1,T}$, where $Y_{1,T}$ is distributed according to

$$
\mathbb{P}(Y_{1,T} = y_1) = \frac{1}{n} \sum_{t=1}^{n} \mathbb{P}(Y_{1,t} = y_1), \quad \forall y_1 \in \mathcal{Y}_1
$$

and (c) comes from the introduction of $X_0 \triangleq X_{0,T}$, $W_1 \triangleq$ $(W_{1,T}, T), W_2 \triangleq W_{2,T}, Y_1 \triangleq Y_{1,T}.$

Next, we show that the introduced auxiliary random variables $W_{1,t} = X_0^{t-1}$, and $W_{2,t} = Y_{1,t+1}^n$ in equation (a) satisfy the Markov chain properties (10).

For $t \in \{1, ..., n\}$:

- $X_{0,t}$ and $W_{1,t} = X_0^{t-1}$ are independent because of the i.i.d. property of the source.
- $U_{1,t} \rightarrow (X_{0,t}, X_t^{t-1}) \rightarrow Y_{1,t+1}^n$, is given by the causal
encoding that the current encoder's output U_t depends encoding, that the current encoder's output $U_{1,t}$ depends

on the future symbols $Y_{1,t+1}^n$ only through the past and current input states $(X_{0,t}, X_0^{t-1})$.

- \bullet $(X_{1,t}, Y_{1,t}) \oplus (X_{0,t}, U_{1,t}) \oplus (X_0^{t-1}, Y_{1,t+1}^n)$ comes from the memoryless property of the channel the memoryless property of the channel.
- $U_{2,t} \leftarrow (X_0^{t-1}, Y_{1,t+1}^n, Y_{1,t}) \leftarrow (X_{0,t}, U_{1,t}, X_{1,t})$ is stated
in Subsection IV Λ below. in Subsection IV-A below.

This implies that the auxiliary random variables X_0 = $X_{0,T}$, $W_1 = (W_{1,T}, T)$, $W_2 = W_{2,T}$, $U_1 = U_{1,T}$, $X_1 =$ $X_{1,T}$, $Y_1 = Y_{1,T}$, $U_2 = U_{2,T}$ satisfy the Markov chains (10). Therefore, the distribution of auxiliary random variables decomposes as in (7).

Now, we reformulate *n*-stage costs with a control design $c \in \mathcal{C}_e(n)$. By defining distributions $\mathbb{P}(U_{1,T}) =$ $(u_1) = \frac{1}{n} \sum_{t=1}^n \mathbb{P}(U_{1,t} = u_1), \quad \forall u_1 \in \mathcal{U}_1 \text{ and } \mathbb{P}(X_{1,T} =$ $x_{1,T}$, $U_{2,T}$ = $u_{2,T}$) = $\frac{1}{n} \sum_{t=1}^{n} \mathbb{P}(X_{1,t} = x_1, U_{2,t}$ = u_2), $\forall (x_1, u_2) \in \mathcal{X}_1 \times \mathcal{U}_2$ and using auxiliary random variables $U_1 = U_{1,T}$, $X_1 = X_{1,T}$, $U_2 = U_{2,T}$, we have

$$
\mathbb{E}[c_P(U_1^n)] = \mathbb{E}\left[\frac{1}{n}\sum_{t=1}^n U_{1,t}^2\right] = \mathbb{E}\left[U_{1,T}^2\right] = \mathbb{E}\left[U_1^2\right],
$$

$$
\mathbb{E}[c_S(X_1^n, U_2^n)] = \mathbb{E}\left[\frac{1}{n}\sum_{t=1}^n (X_{1,t} - U_{2,t})^2\right]
$$

$$
= \mathbb{E}\left[(X_{1,T} - U_{2,T})^2\right] = \mathbb{E}\left[(X_1 - U_2)^2\right].
$$

In conclusion, if a pair of costs $(P, S) \in \mathbb{R}^2$ is achievable, then for all $\varepsilon > 0$, there exists $\bar{n} \in \mathbb{N}$ such that for all $n \geq \bar{n}$, there exists a control design $c \in \mathcal{C}_e(n)$ such that (7) and (8) are satisfied. Moreover, from (5), we also have

$$
\mathbb{E}\Big[|P - c_P(U_1^n)| + |S - c_S(X_1^n, U_2^n)|\Big] \leq \varepsilon
$$

\n
$$
\implies |P - \mathbb{E}[c_P(U_1^n)]| + |S - \mathbb{E}[c_S(X_1^n, U_2^n)]| \leq \varepsilon
$$

\n
$$
\implies |P - \mathbb{E}[U_1^2]| + |S - \mathbb{E}[(X_1 - U_2)^2]| \leq \varepsilon
$$
 (20)

The equation (20) is valid for all $\varepsilon > 0$. This shows that $(X_0, W_1, W_2, U_1, X_1, Y_1, U_2)$ satisfy (7), (8), (9).

A. Sketch Proof of the Markov Chain

$$
\begin{split}\n&\mathcal{P}(U_{2,t}|Y_{1,t},W_{1,t},W_{2,t},X_{0,t},U_{1,t},X_{1,t}) \\
&= \sum_{U_1^{t-1},X_1^{t-1},Y_1^{t-1}} \mathcal{P}(U_1^{t-1}|Y_{1,t},X_0^{t-1},Y_{1,t+1}^n) \\
&\times \quad \mathcal{P}(X_1^{t-1}|Y_{1,t},X_0^{t-1},Y_{1,t+1}^n,U_1^{t-1}) \\
&\times \quad \mathcal{P}(Y_1^{t-1}|Y_{1,t},X_0^{t-1},Y_{1,t+1}^n,U_1^{t-1},X_1^{t-1}) \\
&\times \quad \mathcal{P}(U_{2,t}|Y_{1,t},X_0^{t-1},Y_{1,t+1}^n,U_1^{t-1},X_1^{t-1},Y_1^{t-1}), \\
&= \sum_{U_1^{t-1},X_1^{t-1},Y_1^{t-1}} \mathcal{P}(U_1^{t-1},X_1^{t-1},Y_1^{t-1}U_{2,t}|Y_{1,t},X_0^{t-1},Y_{1,t+1}^n) \\
&= \quad \mathcal{P}(U_{2,t}|Y_{1,t},X_0^{t-1},Y_{1,t+1}^n) \\
&= \quad \mathcal{P}(U_{2,t}|Y_{1,t},W_{1,t},W_{2,t}).\n\end{split}
$$

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