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Quasi-limiting behaviour of the sub-critical Multitype Bisexual Galton-Watson Branching Process

Coralie Fritsch¹, Denis Villemonais^{1,2} and Nicolás Zalduendo³

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Abstract

We investigate the quasi-limiting behaviour of bisexual subcritical Galton-Watson branching processes. While classical subcritical Galton-Watson processes have been extensively analyzed, bisexual Galton-Watson branching processes present unique difficulties because of the lack of the branching property. To prove the existence of and convergence to one or several quasi-stationary distributions, we leverage on recent developments linking bisexual Galton-Watson branching processes extinction to the eigenvalue of a concave operator.

1 Introduction

We study the quasilimiting behaviour of bisexual subcritical Galton-Watson branching processes (bGWbp). This problem has been studied and largely solved decades ago for classical sub-critical Galton-Watson processes (see e.g. [19, 17, 2]), but remains open for its bisexual counterpart. Recent developments in the theory of multi-dimensional bGWbp [9] show that, under appropriate super-additivity assumptions, the extinction of these processes can be characterised by the eigenvalue λ^* associated to a positively homogeneous concave operator and, more precisely, by its position relatively to 1: if $\lambda^* \leq 1$, then the process is eventually extinct almost surely, while, if $\lambda^* > 1$, then, with positive probability, the process grows exponentially fast with rate equal or close to λ^* . We refer the reader to the surveys [1, 12, 16] on bGWbp and [6, 7, 10, 11] for further references. In this paper, we consider the sub-critical case λ^* < 1 and show that the process admits infinitely many quasi-stationary distributions. Under an additional polynomial moment assumption, we prove that the process admits a finitely generated family of quasi-stationary distributions and, under an additional irreducibility assumption, a unique quasi-limiting distribution for compactly supported initial distributions. Although this is similar to the classical Galton-Watson case, λ^* is not in general equal to the exponential survival rate of the process.

Let us introduce more formally our settings and assumptions. Consider $p, q \in \mathbb{N} = \{0, 1, 2, ...\}$ with $p, q \ge 1$, a non-negative vector function $\xi : \mathbb{N}^q \to \mathbb{N}^p$ such that $\xi(0) = 0$, and a family of integrable random vectors $V = (V_{i,\cdot})_{1 \le i \le p}$, taking value in \mathbb{N}^q and whose expectation is denoted by V . We assume that $\sum_{i=1}^{q} V_{i,j} > 0$ for all $j \in \{1, ..., q\}$. We consider a process $Z = (Z_n)_{n \in \mathbb{N}}$ on \mathbb{N}^p which represents the random sequence of the number of couples of each type in the population, and which evolves as follows: given $Z_{n-1} = (Z_{n-1,1},..., Z_{n-1,p})$, we define for *n* ≥ 1 the vector of children of the *n*−th generation ($W_{n,1},...,W_{n,q}$) by

$$
W_{n,j} = \sum_{i=1}^{p} \sum_{k=1}^{Z_{n-1,i}} V_{i,j}^{(k,n)}, \text{ for } 1 \le j \le q,
$$

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where $(V^{(k,n)})_{k,n\in\mathbb{N}}$ is a family of i.i.d. copies of *V*. Then, we set the vector of couples in the *n*−th generation as

$$
Z_n = (Z_{n,1}, \dots, Z_{n,p}) = \xi(W_{n,1}, \dots, W_{n,q}).
$$

The function *ξ* is referred to as the *mating function*. For instance, if *p* = 1, *q* = 2 and if *W_{n,1}* represents the number of females and *Wn*,2 the number of males in the *n*−th generation, classical choices for *ξ* are *ξ*(*x*, *y*) = min{*x*, *y*} (the *perfect fidelity* mating function) and *ξ*(*x*, *y*) = *x* min{1, *y*} (the *promiscuous* mating function). If $p = q = 2$, then $\xi(x, y) = (x, y)$ corresponds to the classical bi-dimensional Galton-Watson process.

For any probability measure μ on \mathbb{N}^p , we use the notation \mathbb{P}_μ and \mathbb{E}_μ for the law and associated expectation of the bisexual Galton-Watson process with initial number of couples *Z*⁰ distributed according to μ . As usual, we use the notations \mathbb{P}_x and \mathbb{E}_x when $\mu = \delta_x$ for some $x \in \mathbb{N}^p$. Note that $\xi(0) = 0$ entails that 0 is absorbing for *Z*.

We are interested in conditions ensuring the existence of a quasi-stationary distribution for *Z*. We recall that a *quasi-stationary distribution* for *Z* is a probability measure v_{QS} on $\mathbb{N}^p \setminus \{0\}$ such that, for all $n \geq 0$, $\mathbb{P}_{\nu_{OS}}(Z_n \neq 0) > 0$ and

$$
\mathbb{P}_{v_{QS}}(Z_n \in \cdot \mid Z_n \neq 0) = v_{QS}(\cdot).
$$

It is known that (we refere the reader to the book [5] and to the surveys [18, 15] for general properties and further examples), for any quasi-stationary distribution *νQS*, there exists *θQS* ∈ (0,1] such that

$$
\mathbb{P}_{v_{QS}}(Z_n \neq 0) = \theta_{QS}^n \text{ and } \mathbb{P}_{v_{QS}}(Z_n \in \cdot, Z_n \neq 0) = \theta_{QS}^n v_{QS}(\cdot).
$$

In what follows, θ_{QS} is referred to as the *absorption parameter* of v_{OS} .

Throughout the paper, we make the assumption that ζ is super-additive and sub-affine:

Assumption (S). There exists $\alpha, \beta \in \mathbb{R}_+^p$ such that

$$
\alpha |x_1 + x_2| + \beta \ge \xi (x_1 + x_2) \ge \xi (x_1) + \xi (x_2), \ \forall x_1, x_2 \in \mathbb{N}^q,
$$
\n(1)

where |·| is the *ℓ* 1 -norm.

This assumption implies that

$$
\mathfrak{M}(z) = \lim_{k \to \infty} \frac{\xi(\lfloor kz \mathbb{V} \rfloor)}{k}, \ \forall z \in \mathbb{R}_+^p
$$

is well defined, bounded over $\mathbb{S} := \{z \in \mathbb{R}^p_+, |z| = 1\}$, positively homogeneous (i.e. $\mathfrak{M}(az) =$ $a\mathfrak{M}(z)$ for all $a > 0$ and all $z \in \mathbb{R}^p_+$) and concave. We refer the reader to Section 3 in [9], where other properties of this functional are derived. In addition, we assume that

Assumption (P). M is primitive, which means that there exists $n_0 \ge 1$ such that $\mathfrak{M}^m(z) > 0$ for all $m \ge n_0$ and $z \in \mathbb{R}_+^p \setminus \{0\}$, where \mathfrak{M}^m is the m −th iterate of \mathfrak{M} .

In particular, this entails that there exists $\lambda^* > 0$ and $z^* \in \mathbb{S}^* := \{z \in (\mathbb{R}_+ \setminus \{0\})^p, |z| = 1\}$ such that the limit

$$
\lim_{n \to \infty} \frac{\mathfrak{M}^n(z)}{(\lambda^*)^n} = \mathcal{P}(z) z^*
$$
 (2)

exists, where $\mathscr{P}:\mathbb{R}^p_+\to\mathbb{R}_+$ is strictly positive on $\mathbb{R}^p_+\setminus\{0\}$, positively homogeneous and concave (see [13] for a general development on this theory and Section 4.1 of [9] for an application to the context of bGWbp). In addition, according to Lemma 37 in [9], $((\lambda^*)^{-n} \mathcal{P}(Z_n))_{n \in \mathbb{N}}$ is a

supermartingale. In particular, λ^* < 1 implies that $(Z_n)_{n\in\mathbb{N}}$ goes extinct almost surely for any initial distribution (see [9] for details).

Our first result states that, in the sub-critical case, the process admits infinitely many quasistationary distributions, indexed by $\theta \in [v_0, 1)$, where

$$
v_0 := \inf \{ v > 0, \text{ such that } \mathbb{E}_z(v^{-T_0}) < +\infty, \ \forall z \in \mathbb{N}^p \}
$$
 (3)

with T_0 the extinction time.

We also define the *exponential convergence parameter* θ_0 by

$$
\theta_0 = \sup_{z \in \mathbb{N}^p \setminus \{0\}} \sup \{ \theta > 0, \liminf_{n \to +\infty} \theta^{-n} \mathbb{P}_z(Z_n \neq 0) > 0 \},\tag{4}
$$

with sup $\emptyset = -\infty$.

Proposition 1.1. *If Assumptions (S) and (P) hold true, then we have* $\theta_0 \le v_0 \le \lambda^*$.

Note that, under the assumptions of Theorem 1.3, we have $v_0 = \theta_0$ (take $f \equiv 1$ in (5)).

Theorem 1.2. *If Assumptions (S) and (P) hold true and λ* [∗] < 1*, then Z admits a continuum of linearly independent quasi-stationary distributions. More precisely, for any* $\theta \in [v_0, 1)$ *, there exists a quasi-stationary distribution with absorption parameter θ.*

The proofs of Proposition 1.1 and Theorem 1.2 are detailed in Section 2, where we first state and prove a general result ensuring the existence of infinitely many quasi-stationary distributions for Feller processes, and then apply it to the sub-critical bGWbp. We emphasize that in the classical Galton-Watson case, much more precise results are known (we recommend to the interested reader the paper [14], where the authors proves a complete result for Galton-Watson processes and gives, in Section 3, a detailed exposition of the history of the problem, with its link to the identification of the Martin boundary of the classical Galton-Watson branching process). We also refer the reader to [8], where a general existence result for quasi-stationary distributions is obtained by a renewal/compactness argument.

We consider now the problem of existence of a finitely generated set of quasi-stationary distributions satisfying an integrability assumption and of convergence of conditional laws toward a quasi-stationary distribution. Since $\mathfrak M$ is concave on \mathbb{R}^p_+ , it is locally Lipschitz on $\mathbb S^*$. We add the following regularity assumption on \mathfrak{M} , where we use the notation, for all $I \subset \{1, ..., p\}$,

 $\mathcal{Z}_I := \{x \in \mathbb{S} \text{ such that } x_i > 0 \ \forall i \in I, x_i = 0 \ \forall i \notin I\},\$

where, here and across the paper, x_i denotes the i^{th} coordinate of x .

Assumption (C). For all $I \neq \emptyset$, \mathfrak{M} is uniformly continuous over \mathcal{Z}_I .

We also consider the following moment condition, based on the exponential convergence parameter θ_0 .

Assumption (M). There exists $\varsigma > 1$ such that $(\lambda^*)^{\varsigma} < \theta_0$ and such that $\mathbb{E}(V^{\varsigma}_i)$ $\binom{r}{i,j}$ < + ∞ for all *i*, *j*.

We emphasize that Assumption (M) implies λ^* < 1, so that the following results applies in the setting of the sub-critical bGWbp. This assumption also requires implicitly that $\theta_0 > 0$, which is the case in many situations.

The proof of the following theorem is detailed in Section 4. It is based on [3] and make use of the following function and measure spaces. Given a positive function φ on $\mathbb{N}^p \setminus \{0\}$, we set

 $L^{\infty}(\varphi) := \{f : \mathbb{N}^p \setminus \{0\} \to \mathbb{R}, \ \|f/\varphi\|_{\infty} < +\infty\},\$ endowed with the norm $\|f\|_{\varphi} = \|f/\varphi\|_{\infty}$

and

 $\mathcal{M}(\varphi) := \{ \mu \text{ non-negative measure on } \mathbb{N}^p \setminus \{0\} \text{ such that } \mu(\varphi) < +\infty \}$

where $\mu(\varphi) = \int_{\mathbb{N}^p \setminus \{0\}} \varphi(z) \mu(\mathrm{d}z)$.

Finally, when Assumption (M) is enforced, we define, for all $a \in (1, \zeta)$ such that $(\lambda^*)^a < \theta_0$, the function $Q_a : \mathbb{N}^p \setminus \{0\} \longrightarrow \mathbb{R}_+$ by

$$
Q_a(z) = \mathcal{P}(z)^a.
$$

Note that $\inf Q_a > 0$.

In the following theorem, we say that a process *Z* is aperiodic if, for all $z \in \mathbb{N}^p$, either $\mathbb{P}_z(\exists n \geq 0)$ $1, Z_n = z$ = 0 or there exists $n_0(z) \ge 0$ such that, $\forall n \ge n_0(z)$, $\mathbb{P}_z(Z_n = z) > 0$.

Theorem 1.3. *We assume that Assumptions (S), (P), (C) and (M) are in place and that the process Z* is aperiodic. We fix $a \in (1, \varsigma)$ such that $(\lambda^*)^a < \theta_0$.

Then there exist $\ell \geq 1$ *and a family of quasi-stationary distributions* $v_1, \ldots, v_\ell \in \mathcal{M}(Q_a)$ *with absorption parameter* θ_0 *for Z such that any quasi-stationary distribution* $v \in \mathcal{M}(Q_a)$ *with absorption parameter* θ_0 *for Z is a convex combination of* v_1, \ldots, v_ℓ *. In addition, there exists a bounded function* $j : \mathbb{N}^p \setminus \{0\} \to \mathbb{N}$ and there exists, for each $i \in \{1, ..., \ell\}$, a non-negative non*identically zero function* $\eta_i \in L^\infty(Q_a)$ *such that, for all* $f \in L^\infty(Q_a)$ *, all* $n \ge 1$ *and all* $z \in \mathbb{N}^p \setminus \{0\}$ *,*

$$
\left| \theta_0^{-n} n^{-j(z)} \mathbb{E}_z(f(Z_n) \mathbf{1}_{Z_n \neq 0}) - \sum_{i=1}^{\ell} \eta_i(z) \nu_i(f) \right| \leq \alpha_n Q_a(z) \|f\|_{Q_a},\tag{5}
$$

where α_n goes to 0 *when* $n \to +\infty$ *which does not depend on* f nor on *z* (but may depend on *a*).

In addition, the set $E := \{z \in \mathbb{N}^p \setminus \{0\}, \sum_{i=1}^\ell \eta_i(z) = 0\}$ *is finite, and any quasi-stationary distributions for Z in* $\mathcal{M}(Q_a)$ *with absorption parameter strictly smaller than* θ_0 *is supported by* E.

Finally, there is no quasi-stationary distribution for Z in M(*Qa*) *with absorption parameter strictly larger than* θ_0 *and, for all* $z \in \mathbb{N}^p \setminus \{0\}$ *, the conditional distribution* $\mathbb{P}_z(Z_n \in \cdot \mid Z_n \neq 0)$ *converges in* $\mathcal{M}(Q_a)$ *.*

Under the additional assumption that the process *Z* is irreducible (meaning that $∀x, γ ∈$ $\mathbb{N}^p \setminus \{0\}$, there exists $n \geq 0$ such that $\mathbb{P}_x(Z_n = y) > 0$, one obtains that, for all $z \in \mathbb{N}^p \setminus \{0\}$,

$$
\theta_0 = \sup \{ \theta > 0, \, \liminf_{n \to +\infty} \theta^{-n} \mathbb{P}_z(Z_n \neq 0) > 0 \},\,
$$

does not depend on *z*. In the irreducible case, we have the following result, proved in Section 5.

Theorem 1.4. *Assume that Assumptions (S), (P), (C) and (M) hold true and that the process Z is aperiodic and irreducible. We fix* $a \in (1, \varsigma)$ *such that* $(\lambda^*)^a < \theta_0$.

Then the process Z admits a unique quasi-stationary distribution v_{OS} *in* $\mathcal{M}(Q_a)$ *. Its absorption parameter* θ_{QS} *is equal to* θ_0 *and there exists a unique positive function* η_{QS} *in* $L^\infty(Q_a)$ *, such that*

$$
\mathbb{P}_{\nu_{QS}}(Z_n \neq 0) = \theta_0^n \quad \text{and} \quad \mathbb{E}_z(\eta_{QS}(Z_n) \mathbf{1}_{Z_n \neq 0}) = \theta_0^n \eta_{QS}(z), \quad \forall n \ge 0, \ z \in \mathbb{N}^p \setminus \{0\} \tag{6}
$$

and η_{OS} *is lower bounded away from* 0*. In addition, there exist* $C > 0$ *and* $\gamma \in (0,1)$ *such that, for* any probability measure $\mu \in \mathcal{M}(Q_a)$ and any function $f \in L^{\infty}(Q_a)$, we have

$$
\left|\theta_0^{-n}\mathbb{E}_{\mu}\left(f(Z_n)\mathbf{1}_{Z_n\neq 0}\right) - \mu(\eta_{QS})\nu_{QS}(f)\right| \le C\gamma^n \mu(Q_a) \|f\|_{Q_a}, \quad \forall n \ge 0
$$
\n(7)

and

$$
\left|\mathbb{E}_{\mu}\left(f(Z_n) \mid Z_n \neq 0\right) - \nu_{QS}(f)\right| \le C\gamma^n \mu(Q_a) \left\|f\right\|_{Q_a}, \quad \forall n \ge 0. \tag{8}
$$

Finally, for all probability measure $\mu \in \mathcal{M}(\eta_{OS})$ *,*

$$
\mathbb{P}_{\mu}(Z_n \in \cdot \mid Z_n \neq 0) \xrightarrow[n \to +\infty]{TV} V_{QS},\tag{9}
$$

where T V refers to the total variation convergence.

Theorem 1.4 states the existence and the uniqueness of a quasi-stationary distribution *νQS* integrating the Lyapunov function Q_a . Moreover, (8) and (9) establish the convergence of the law of the process conditioning to the non-extinction towards *νQS* for all initial distribution *µ* integrating Q_a or the eigenfunction η_{QS} , with moreover an exponential speed of convergence for μ integrating Q_a . In particular, the convergence towards v_{OS} holds for any Dirac mass μ = δ _{*z*}, with $z \in E$, meaning that v_{OS} is a Yaglom limit. In addition, (7) gives the convergence of the law of the process without conditioning. In particular, taking $f \equiv 1$ in (7) leads to the speed of convergence of the extinction probability towards 0 at speed exactly θ_0 .

The proofs of the last two results rely in part on recent advances in the theory of quasistationary distributions, which allow to derive several properties from Foster-Lyapunov type criteria, given in our case by the following proposition.

Proposition 1.5. *If Assumptions (S), (P), (C) and (M) hold true, then, for any* $a \in (1, \varsigma)$ *such that* $(\lambda^*)^a < \theta_0$, we have

$$
\mathbb{E}(Q_a(Z_1)\mathbf{1}_{Z_1\neq 0}|Z_0=z) \leq \theta_a Q_a(z) + C_a, \forall z \in \mathbb{N}^p \setminus \{0\}
$$
\n(10)

for some constants $\theta_a \in [0, \theta_0)$ *and* $C_a > 0$ *.*

Remark 1.1. We already observed that $\theta_0 \leq \lambda^*$. One also checks that, at least under the assumptions of Theorem 1.4, if $\mathscr P$ is not linear, then $\theta_0 < \lambda^*$. Indeed, under these assumptions, let v_{QS} be the quasi-stationary distribution from Theorem 1.4, so that

$$
\theta_0^n v_{QS}(\mathscr{P}) = \sum_{z \in \mathbb{N}^p \setminus \{0\}} v_{QS}(\{z\}) \mathbb{E}_z(\mathscr{P}(Z_n)).
$$

Since under the assumptions of Theorem 1.4, the process is irreducible, the support of *νQS* is $\mathbb{N}^p \setminus \{0\}$ and hence $\mathscr P$ is concave not affine on the support of v_{QS} . Using (8), we deduce that there exist $z_0 \in \mathbb{N}^p \setminus \{0\}$ and $n \ge 1$ such that $\mathcal P$ is concave not affine on the support of $\mathbb P_{z_0}(Z_n \in \cdot)$, so that, by Jensen (strict) inequality,

$$
\mathbb{E}_{z_0}\left(\mathcal{P}(Z_n)\right) < \mathcal{P}(\mathbb{E}_{z_0}(Z_n)) \leq \mathcal{P}(\mathfrak{M}^n(z_0)) = (\lambda^*)^n \mathcal{P}(z_0),
$$

where we used $\mathbb{E}_{z_0}(Z_n) \leq \mathfrak{M}^n(z_0)$ (see Lemma 27 in [9]) and the fact that \mathscr{P} is superadditive (since M is superadditive by Assumption (S) and by definition of $\mathscr P$ given by (2)) for the last inequality, and the definition of $\mathscr P$ (see (2)) for the last equality. Similarly, we have $\mathbb E_z(\mathscr P(Z_n)) \leq$ $(\lambda^*)^n \mathscr{P}(z)$ for all $z \in \mathbb{N}^p \setminus \{0\}$. We conclude that

$$
\theta_0^n v_{QS}(\mathcal{P}) < \sum_{z \in \mathbb{N}^p \backslash \{0\}} v_{QS}(\{z\}) (\lambda^*)^n \mathcal{P}(z) = (\lambda^*)^n v_{QS}(\mathcal{P}).
$$

This implies that $\theta_0 < \lambda^*$ when $\mathscr P$ is not linear.

It is also known that $\mathfrak{M}(z) = \sup_{k \geq 1} \frac{1}{k}$ $\frac{1}{k} \mathbb{E}_{kz}(Z_1)$ (see Corollary 7 in [9]), and the same calculation as above shows that, if $\mathbb{E}_{z_0}(Z_1) < \mathfrak{M}(z_0)$ for some $z_0 \in \mathbb{N}^p$, then $\theta_0 < \lambda^*$ (even if \mathscr{P} is linear). *Remark* 1.2. When $\theta_0 = \lambda^*$, Assumption (M) requires polynomial moments for *V*, with exponents that can be arbitrarily close to 1. We leave as an open question whether, in this case, *L* log*L* type criteria can be obtained, as it is the case in the classical Galton-Watson case (for which $\theta_0 = \lambda^*$).

In the situation where $\theta_0 < \lambda^*$, one can not choose *ς* arbitrarily close to 1 in Assumption (M), since it imposes $\zeta > \frac{|\log \theta_0|}{|\log \lambda^*|}$ $\frac{|\log \theta_0|}{|\log \lambda^*|}$. We leave as an open question whether $\mathbb{E}(V_i^{\varsigma})$ $\binom{r\varsigma}{i,j}$ < + ∞ for some $\zeta \le \frac{|\log \theta_0|}{|\log \lambda^*|}$ may be sufficient to obtain the conclusion of Theorem 1.3 or 1.4, at least in some particular cases.

In Section 2, we state a general result for the existence of infinitely many quasi-stationary distributions, then we prove Proposition 1.1 and Theorem 1.2. Section 3 is devoted to the proof of Proposition 1.5. Theorems 1.3 and 1.4 are proved in Sections 4 and 5 respectively. For the proofs of the last two theorems, we make use of Assumption (E) of [4] which is recalled in Appendix. In Appendix we also state a general result implying Assumption (E) of [4].

2 A simple criterion for the existence of infinitely many quasi-stationary distributions and application to the bGWbp

In Section 2.1, we first state and prove a result on the existence of a continuum of quasi-stationary distributions for general sub-Markov kernels on Polish spaces. In Section 2, we apply this result to the bGWbp.

2.1 General setting

Let (X, d) be a Polish space endowed with a σ -algebra $\mathscr X$ that contains the toplogy induced by *d*. Let *K* be a sub-Markov kernel on *X*, which means that *K* is a non-negative kernel on *X* such that $K(x, X) \leq 1$ for all $x \in X$. We assume that *K* has Feller regularity, meaning that the associated operator $K: f \mapsto Kf := \int_X K(\cdot, dy) f(y)$ is a well defined functional from $C_b(X)$ to $C_b(X)$ (the set of continuous bounded functions on *X*). We say that a probability measure μ on *X* is a quasi-stationary distribution for *K* if, for some $\lambda \in (0,1]$,

$$
\int_X \mu(\mathrm{d} x) K(x,\cdot) = \lambda \mu(\cdot).
$$

Our aim is to prove that *K* admits infinitely many quasi-stationary distributions under mild assumptions.

In the following theorem, we say that a sequence $(x_n)_{n\in\mathbb{N}}\in X^{\mathbb{N}}$ tends to infinity if it eventually leaves any compact set: for all compact subset *L* of *X*, there exists $n_L \geq 0$ such that, for all *n* ≥ *n*_L, *x*_{*n*} ∉ *L*. Similarly, we say that a real functional on *X* tends to infinity when $x \to +\infty$ if, for all $R > 0$, there exists a compact subset L_R of X such that the functional is bounded below by *R* on $X \setminus L_R$. We also define the iterated kernels K^n as usual by

$$
K^{0}(x, A) = \mathbf{1}_{x \in A}, K^{\ell+1}(x, A) = \int_{X} K^{\ell}(x, dy) K(y, A) \,\forall x \in X, A \in \mathcal{X}, \ell \ge 0.
$$

Theorem 2.1. Assume that there exists a > 0 and $(x_n)_{n\in\mathbb{N}}$ \in $X^{\mathbb{N}}$ tending to infinity such that

$$
\sum_{\ell=0}^{\infty} a^{-\ell} K^{\ell}(x_n, X) < +\infty, \ \forall \, n \ge 0 \tag{11}
$$

and let

 $a_0 := \inf\{a > 0, \text{ s.t. there exists } (x_n)_{n \in \mathbb{N}} \in X^{\mathbb{N}} \text{ tending to infinity such that (11) holds true.} \}$ (12)

Assume also that there exists $a_1 > a_0$ *such that*

$$
\sum_{\ell=0}^{\infty} a_1^{-\ell} K^{\ell}(x, X) \xrightarrow[x \to +\infty]{} +\infty.
$$
 (13)

Then K admits an infinite family of quasi-stationary distributions $(\mu_{\lambda})_{\lambda \in [a_0, a_1)}$, with $\int_X \mu_{\lambda}(\mathrm{d} x) K(x, \cdot) =$ $\lambda \mu_{\lambda}$ *for all* $\lambda \in [a_0, a_1)$ *.*

Proof of Theorem 2.1. Fix $\lambda \in (a_0, a_1)$. Since $\lambda > a_0$ then there exists $(x_n)_{n \in \mathbb{N}} \in X^{\mathbb{N}}$ tending to infinity such that, for all $n \geq 0$, the probability measure

$$
\mu_{\lambda,n}=\frac{\sum_{\ell=0}^\infty \lambda^{-\ell}K^\ell(x_n,\cdot)}{\sum_{\ell=0}^\infty \lambda^{-\ell}K^\ell(x_n,X)}
$$

is well defined. We have, for all $n \geq 0$,

$$
\mu_{\lambda,n}K(\cdot) := \int_X \mu_{\lambda,n}(\mathrm{d}x)K(x,\cdot) = \frac{\sum_{\ell=0}^{\infty} \lambda^{-\ell} K^{\ell+1}(x_n,\cdot)}{\sum_{\ell=0}^{\infty} \lambda^{-\ell} K^{\ell}(x_n, X)}
$$

$$
= \frac{\lambda \sum_{\ell=0}^{\infty} \lambda^{-(\ell+1)} K^{\ell+1}(x_n,\cdot)}{\sum_{\ell=0}^{\infty} \lambda^{-\ell} K^{\ell}(x_n, X)}
$$

$$
= \frac{\lambda \sum_{\ell=0}^{\infty} \lambda^{-\ell} K^{\ell}(x_n,\cdot) - \lambda \delta_{x_n}(\cdot)}{\sum_{\ell=0}^{\infty} \lambda^{-\ell} K^{\ell}(x_n, X)}
$$
(14)
$$
\leq \lambda \mu_{\lambda,n}(\cdot).
$$

We deduce that, for all $\ell \geq 1$,

$$
\mu_{\lambda,n}K^{\ell}\leq \lambda^{\ell}\mu_{\lambda,n}.
$$

In particular, we obtain that

$$
\int_X \mu_{\lambda,n}(\mathrm{d}x) \sum_{\ell=0}^\infty a_1^{-\ell} K^{\ell}(x,X) = \sum_{\ell=0}^\infty a_1^{-\ell} \mu_{\lambda,n} K^{\ell}(X) \le \sum_{\ell=0}^\infty \left(\frac{\lambda}{a_1}\right)^{\ell} = \frac{a_1}{a_1 - \lambda}.\tag{15}
$$

Since we assumed that $\sum_{\ell=0}^{\infty} a_1^{-\ell} K^{\ell}(x,X)$ goes to infinity when $x \to \infty$, this and Prokhorov's Theorem imply that the sequence of measures $(\mu_{\lambda,n})_{n\in\mathbb{N}}$ is relatively compact for the weak convergence topology.

Let μ_{λ} be any adherent point of $(\mu_{\lambda,n})_{n \in \mathbb{N}}$, and observe that μ_{λ} is a probability measure. Since *K* is assumed to preserve bounded continuous functions, we deduce that, up to a subsequence and for the weak convergence topology,

$$
\mu_{\lambda} K = \lim_{n \to +\infty} \mu_{\lambda, n} K.
$$

But we computed (see (14))

$$
\mu_{\lambda,n}K = \lambda \mu_{\lambda,n} - \frac{\lambda \delta_{x_n}}{\sum_{\ell=0}^{\infty} \lambda^{-\ell} K^{\ell}(x_n, X)},
$$

where, by assumption,

$$
\sum_{\ell=0}^{\infty} \lambda^{-\ell} K^{\ell}(x_n, X) \ge \sum_{\ell=0}^{\infty} a_1^{-\ell} K^{\ell}(x_n, X) \xrightarrow[n \to +\infty]{} +\infty.
$$

We deduce that $\lim_{n\to+\infty} \mu_{\lambda,n} K = \lim_{n\to+\infty} \lambda \mu_{\lambda,n}$, which is equal to $\lambda \mu_\lambda$. We thus proved that $\mu_{\lambda} K = \lambda \mu_{\lambda}$, which means that μ_{λ} is a quasi-stationary distribution for *K* with absorption parameter *λ*.

It remains to prove that there exists a quasi-stationary distribution with absorption parameter a_0 . Since the set of probability measures μ_{λ} , $\lambda \in (a_0, (a_0 + a_1)/2)$, satisfies (15) with a uniform upper bound, we deduce that it is relatively compact. One checks as above that any limit point, when $\lambda \to a_0$, of this family is a quasi-stationary distribution for *K* with absorption parameter a_0 . This concludes the proof of Theorem 2.1. \Box

Remark 2.1*.* One easily checks from the proof that the condition that *K* preserves bounded continuous functions is stronger than necessary. Actually, it would be sufficient to assume that *K* sends a generating family of bounded continuous functions into a family of bounded continuous functions. For instance, it sufficient to assume that *K* sends compactly supported continuous functions to bounded continuous functions.

Remark 2.2*.* One can easily extends this result to the general setting of non-conservative kernels *K*. More precisely, assume that there exists a positive function *h* on *X* such that $Kh \leq ch$ for some *c* > 0. Then the kernel defined by

$$
\widetilde{K}(x, A) = \frac{1}{ch(x)} \int_{A} K(x, dy) h(y), \ \forall x \in X, \ \forall A \in \mathcal{X},
$$

is a sub-Markov kernel to which Theorem 2.1 may apply. Note that the assumptions would translate into: for any continuous function $f: X \to \mathbb{R}$, $x \in X \mapsto \frac{1}{h(x)} \int_X K(x, dy) h(y) f(y)$ is continuous; for all $n \ge 0$, $\sum_{\ell=0}^{\infty} a_0^{-\ell} \int_X K^{\ell}(x_n, dy) h(y) < +\infty$ and $\frac{1}{h(x)} \sum_{\ell=0}^{\infty} a_1^{-\ell} \int_X K^{\ell}(x, dy) h(y) \xrightarrow[x \to +\infty]{}$ +∞. In addition, the resulting quasi-stationary measures may not be a finite measure if *h* is not lower bounded (we only obtain that $\int_X \mu_\lambda(dy) h(y) < +\infty$ for all $\lambda \in [a_0, a_1)$). On the other hand, if *h* is lower bounded away from 0, then μ_{λ} is a finite measure.

2.2 Proof of Proposition 1.1 and Theorem 1.2

In the context of bGWbp, we choose $K(x,dy) = \mathbb{P}_x(Z_1 \in dy, Z_1 \neq 0)$ for all $x \in X = \mathbb{N}^p \setminus \{0\}$. We first prove Proposition 1.1, as well as that $a_0 \le v_0$, where a_0 is defined in Theorem 2.1 and v_0 is defined in (3). Then we prove Theorem 1.2.

First, from their definitions, it is clear that $v_0 \le 1$ and $\theta_0 \le 1$. For all $\theta < \theta_0$, there exist *v* ∈ (θ , 1) and *z* ∈ $\mathbb{N}^p \setminus \{0\}$ such that liminf_{*n*→∞} $v^n \mathbb{P}_z(Z_n \neq 0) > 0$. For such *v* and *z*, we have $\mathbb{E}_z(v^{-T_0}) = \sum_{\ell=0}^{\infty} v^{-\ell} \mathbb{P}_z(T_0 = \ell)$ with

$$
\liminf_{\ell \to \infty} v^{-\ell} \mathbb{P}_z(T_0 = \ell) = \liminf_{\ell \to \infty} v^{-\ell} \left(\mathbb{P}_z(Z_{\ell-1} \neq 0) - \mathbb{P}_z(Z_{\ell} \neq 0) \right)
$$

$$
= (v^{-1} - 1) \liminf_{\ell \to \infty} v^{-\ell} \mathbb{P}_z(Z_{\ell} \neq 0) > 0.
$$

Then $\mathbb{E}_z(v^{-T_0}) = \infty$ and hence $v \ge v_0$. We deduce that $v_0 \ge \theta$ and hence $v_0 \ge \theta_0$.

In order to prove that $a_0 \le v_0$, we first observe that, for all $a > v > v_0$, for all $z \in \mathbb{N}^p \setminus \{0\}$ and

$$
\sum_{\ell \ge 0} a^{-\ell} K^{\ell}(z, X) = \sum_{\ell \ge 0} a^{-\ell} \mathbb{P}_z(T_0 \ge \ell + 1) \le \sum_{\ell \ge 0} a^{-\ell} \nu^{\ell + 1} \mathbb{E}_z(\nu^{-T_0}) \le C_{a, \nu} \mathbb{E}_z(\nu^{-T_0}) < \infty.
$$

This shows that $a_0 \le v_0$.

Using the fact that $\mathscr P$ is lower bounded away from 0 on $\mathbb N^p\setminus\{0\}$ (as $\mathscr P$ is strictly positive and positively homogeneous) and that *K* vanishes at 0, we deduce that there exists a constant *c* > 0 such that, for all $\ell \geq 1$ and all $x \in \mathbb{N}^p$,

$$
K^{\ell}(x,X) \leq \frac{1}{c} \mathbb{E}_x(\mathscr{P}(Z_{\ell}))\,.
$$

In addition, $(\lambda^*)^{-n} \mathscr{P}(Z_n)$ is a supermartingale (see Lemma 37 in [9]), so that, for any $a \in (\lambda^*, 1)$,

$$
\sum_{\ell=0}^{\infty} a^{-\ell} K^{\ell}(x,X) \leq \frac{1}{c} \sum_{\ell=0}^{\infty} a^{-\ell} (\lambda^*)^{\ell} \mathcal{P}(x) < +\infty.
$$

This shows that $v_0 \leq \lambda^*$, then Proposition 1.1 holds.

Since \mathfrak{M} is primitive by Assumption (P), Theorem 6 in [9] implies that, for all $z \in \mathbb{N}^p \setminus \{0\}$, there exists $c_0 > 0$ and $k_0 \ge 1$ such that for all $k \ge k_0$

$$
\mathbb{P}(Z_{n_0} \ge c_0(1,\ldots,1) \mid Z_0 = kz) > 0.
$$

Applying this result for z taken in each element of the canonical basis of \mathbb{N}^p and using the super-additivity of ξ , one deduces that there exist uniform constants $r > 0$ and $c'_0 > 0$ such that

$$
\inf_{|z|>r} \mathbb{P}(Z_{n_0} \ge c'_0(1,\ldots,1) \mid Z_0=z) > 0.
$$

Using the Markov property and by iteration, this implies that, for all $n \geq 1$,

$$
\mathbb{P}_z(Z_n=0) \xrightarrow[|z|\to+\infty]{} 0.
$$

In particular, we deduce from Fatou's Lemma that, for any $a_1 \in (0,1)$,

$$
\sum_{\ell=0}^{\infty} a_1^{-\ell} \mathbb{P}_z(Z_{\ell} \neq 0) \xrightarrow[|z| \to +\infty]{} +\infty.
$$

and hence (13) holds true for any $a_1 \in (a_0, 1) \supset (v_0, 1)$.

We deduce that Theorem 2.1 applies to the subcritical bGWbp. Since quasi-stationary distributions with different absorption parameters are linearly independent, this proves Theorem 1.2.

3 Proof of Proposition 1.5

We start with the following technical lemma on the regularity of \mathcal{P} .

Lemma 3.1. *Assume that Assumptions (S), (P), (C) and (M) hold true. Then, for all I* ⊂ {1,..., p }*,* $I \neq \emptyset$, the operator P restricted to \mathcal{I}_I is uniformly continuous.

Proof of Lemma 3.1. Since $\mathscr{P} \circ \mathfrak{M}^{n_0} = (\lambda^*)^{n_0} \mathscr{P}$, it is sufficient to prove that $\mathscr{P} \circ \mathfrak{M}^{n_0}$ is uniformly continuous on \mathcal{Z}_I . But $\mathfrak{M}^{n_0}(\mathcal{Z}_I) \subset \mathfrak{M}^{n_0}(\mathbb{S})$ which is relatively compact in $(\mathbb{R}_+ \setminus \{0\})^p$ (since \mathfrak{M} is primitive and bounded by assumption), and \mathfrak{P} is concave and thus locally Lipschitz in $(\mathbb{R}_+ \setminus \{0\})^p$, thus it is sufficient to prove that \mathfrak{M}^{n_0} is uniformly continuous on \mathcal{Z}_I .

Let us first observe that \mathfrak{M} is locally uniformly continuous on $(\varepsilon + \mathbb{R}_+ \setminus \{0\}) \mathcal{Z}_I$ for any $\varepsilon > 0$, since $\mathfrak M$ is uniformly continuous on $\mathcal Z_I$, bounded on $\mathbb S \supset \mathcal Z_I$, and positively homogeneous. In order to conclude, it remains to prove that there exist $J \subset \{1, ..., p\}$ and $\varepsilon > 0$ such that $\mathfrak{M}(\mathcal{Z}_I) \subset (\varepsilon + \mathbb{R}_+ \setminus \{0\}) \mathcal{Z}_I$ (the result then follows by induction). Let $x \in \mathcal{Z}_I$ and let $J = \{i \in I\}$ $\{1,\ldots,p\}$, $\mathfrak{M}(x)_i > 0\}$. Then, for all $y \in \mathcal{Z}_I$, there exists $\delta > 0$ such that $\delta x \leq y$, so that, using the fact that M is increasing and homogeneous, $\mathfrak{M}(\gamma) \geq \delta \mathfrak{M}(x)$ and hence $J \subset \{i \in \{1, ..., p\}, \mathfrak{M}(\gamma)\}\geq \delta \mathfrak{M}(x)$ 0}. Reciprocally, $\{i \in \{1, ..., p\}, \mathfrak{M}(y)_i > 0\} \subset J$, so that $\mathfrak{M}(y) \in \mathbb{R}_+ \mathcal{Z}_J$. Since on $\mathbb S$, the norm of $\mathfrak M$ is bounded from below away from 0 and from above, we have $\mathfrak{M}(\mathcal{Z}_I) \subset [\varepsilon_1, \varepsilon_2] \mathcal{Z}_J$ for some ε_1 , ε_2 > 0. This implies that ${\mathfrak M}^2$ is uniformly continuous on ${\mathcal{Z}}_I.$ We conclude by iteration. \Box

Let us now proceed with the proof of Proposition 1.5 in two steps: first we show that, for all $z \in \mathbb{R}^p_+ \setminus \{0\},\$

$$
\limsup_{n \to +\infty} \frac{1}{|n|^a} \mathbb{E} \left[Q_a(Z_1) \mathbf{1}_{Z_1 \neq 0} \mid Z_0 = \lfloor nz \rfloor \right] = (\lambda^*)^a Q_a(z); \tag{16}
$$

Second, we extend this to a uniform convergence in *z* by using the regularity of \mathcal{P} . *Step 1.* As $\mathscr P$ is positively homogeneous, we have for any $z \in \mathbb{R}_+^p \setminus \{0\}$,

$$
\mathbb{E}\left[Q_a(Z_1)\mathbf{1}_{Z_1\neq0} \mid Z_0 = \lfloor nz \rfloor\right] = \mathbb{E}\left[\left(\mathcal{P}\circ\xi\left(\sum_{i=1}^p\sum_{k=1}^{\lfloor nz_i\rfloor}V_{i,\cdot}^{(k,1)}\right)\right)^a\right]
$$

$$
= |n|^a \mathbb{E}\left[\left(\mathcal{P}\left(\frac{1}{n}\xi\left(\sum_{i=1}^p\sum_{k=1}^{\lfloor nz_i\rfloor}V_{i,\cdot}^{(k,1)}\right)\right)\right)^a\right].
$$

According to Lemma 30 in [9], we have

$$
\frac{1}{n} \xi \left(\sum_{i=1}^p \sum_{k=1}^{\lfloor n z_i \rfloor} V_{i,\cdot}^{(k,1)} \right) \xrightarrow[n \to +\infty]{a.s.} \mathfrak{M}(z).
$$

Let $I \subset \{1, ..., p\}$ be such that $\mathfrak{M}(z) \in \mathbb{R}_+^* \mathcal{Z}_I$, then the above convergence entails that, almost surely, for *n* large enough, $\frac{1}{n} \xi \left(\sum_{i=1}^p \sum_{k=1}^{\lfloor nz_i \rfloor} V_{i,\cdot}^{(k,1)} \right) \in \mathbb{R}_+^* \mathcal{Z}_J$ for some $J \supset I$. Using Corollary 7 in [9], we have, for all $n \ge 1$,

$$
\mathfrak{M}(z) \geq \frac{1}{n} \mathbb{E} [Z_1 | Z_0 = \lfloor nz \rfloor] = \frac{1}{n} \mathbb{E} \left[\xi \left(\sum_{i=1}^p \sum_{k=1}^{\lfloor nz_i \rfloor} V_{i,\cdot}^{(k,1)} \right) \right],
$$

so that we also have $J \subset I$ almost surely.

Since $\mathscr P$ restricted to $\mathbb R_+^* \mathcal Z_I$ is continuous since it is concave, and since $\mathfrak M \circ \mathscr P = \lambda^* \mathscr P$ by (2), we deduce that

$$
\mathscr{P}\left(\frac{1}{n}\xi\left(\sum_{i=1}^p\sum_{k=1}^{\lfloor n z_i\rfloor}V_{i,\cdot}^{(k,1)}\right)\right)\xrightarrow[n\to+\infty]{a.s}\mathscr{P}\left(\mathfrak{M}(z)\right)=\lambda^*\mathscr{P}(z).
$$

In addition, since ξ is sub-affine, since $\frac{1}{n}\sum_{i=1}^p\sum_{k=1}^{\lfloor nz_i\rfloor}V^{(k,1)}_{i,\cdot}$ is bounded in L^ς by Assumption (M) and $\mathscr P$ being positively homogeneous, we deduce that $\mathscr P\big(\frac{1}{n}\big)$ $\frac{1}{n} \xi \left(\sum_{i=1}^p \sum_{k=1}^{\lfloor nz_i \rfloor} V^{(k,1)}_{i,\cdot} \right)$ is also bounded in L^{ς} (where $\varsigma > a$), and we deduce from the dominated convergence theorem that

$$
\mathbb{E}\left[\left(\mathscr{P}\left(\frac{1}{n}\xi\left(\sum_{i=1}^p\sum_{k=1}^{\lfloor nz_i\rfloor}V_{i,\cdot}^{(k,1)}\right)\right)\right)^a\right]\xrightarrow[n\to+\infty]{}(\lambda^*)^a\mathscr{P}(z)^a.
$$

We deduce that (16) holds true. This concludes the first step of the proof.

Step 2. Fix a non-empty $I \subset \{1, ..., p\}$. Let $\varepsilon_1 > 0$ be a small number that will be fixed later and let $z^1, \ldots, z^{m_1} \in (1+\varepsilon_1)\mathcal{Z}_I$, $m_1 = m(\varepsilon_1) \ge 1$, be such that, for all $z \in \mathcal{Z}_I$, there exists $k = k(z, \varepsilon_1) \in$ $\{1,\ldots,m_1\}$ such that $z \leq z^k$ and $|z-z^k| \leq 2\varepsilon_1$. The existence of this finite family z^1,\ldots,z^{m_1} easily follows from a compactness argument. Let $n_1 = n_1(\varepsilon_1)$ be large enough so that, for all $n \ge n_1$, for all $k \in \{1, ..., m_1\}$,

$$
\frac{1}{|n|^a} \mathbb{E}\left[Q_a(Z_1)\mathbf{1}_{Z_1\neq 0} \mid Z_0 = \lfloor nz^k \rfloor\right] \leq (1+\varepsilon_1)(\lambda^*)^a \mathcal{P}(z^k)^a.
$$

Then, since Q_a is non-decreasing and since Z_1 (with $Z_0 = \lfloor nz \rfloor$) is stochastically non-decreasing with respect to *z*, we deduce that, for all $n \geq n_1$ and all $z \in \mathcal{Z}_I$,

$$
\frac{1}{|n|^a} \mathbb{E} \left[Q_a(Z_1) \mathbf{1}_{Z_1 \neq 0} \mid Z_0 = \lfloor nz \rfloor \right] \le (1 + \varepsilon_1)(\lambda^*)^a \mathcal{P}(z^{k(z, \varepsilon_1)})^a
$$

$$
\le (1 + \varepsilon_1)(\lambda^*)^a \mathcal{P}(z)^a \left(\sup_{x, y \in \mathcal{Z}_I, |(1 + \varepsilon_1)x - y| \le 2\varepsilon_1} \frac{\mathcal{P}((1 + \varepsilon_1)x)}{\mathcal{P}(y)} \right)^a.
$$
 (17)

Since ${\mathscr P}$ is uniformly continuous on ${\mathcal{Z}}_I$ and lower bounded away from 0 on ${\mathcal{Z}}_I$, we deduce that there exists ε_1 small enough so that

$$
\theta_a:=(1+\varepsilon_1)^{a+1}(\lambda^*)^a\left(\sup_{x,y\in \mathcal{Z}_I, |(1+\varepsilon_1)x-y|\leq 2\varepsilon_1}\frac{\mathcal{P}(x)}{\mathcal{P}(y)}\right)^a<\theta_0.
$$

This and (17) implies that

$$
\frac{1}{|n|^a}\mathbb{E}\left[Q_a(Z_1)\mathbf{1}_{Z_1\neq0} \mid Z_0=\lfloor nz\rfloor\right]\leq \theta_a\mathcal{P}(z)^a.
$$

Since any $z \in \mathbb{N}^p \setminus \{0\} \cap (\mathbb{R}_+ \setminus \{0\}) \mathcal{Z}_I$ can be written as $n \frac{z}{|z|}$ with $n = |z|$ and $\frac{z}{|z|} \in \mathcal{Z}_I$, we deduce that, for any such *z* with $|z| \ge n_1$,

$$
\frac{1}{|z|^a} \mathbb{E}\left[Q_a(Z_1)\mathbf{1}_{Z_1\neq 0} \mid Z_0 = z\right] \leq \theta_a \mathcal{P}(z/|z|)^a.
$$

Finally, the homogeneity of P allows us to conclude that Proposition 1.5 holds true for all $z \in$ $\mathbb{N}^p\setminus\{0\}\cap(\mathbb{R}_+\setminus\{0\})$ \mathcal{Z}_I . Since there are only finitely many subsets *I* of $\{1,\ldots,p\}$, Proposition 1.5 holds true for all $z \in \mathbb{N}^p \setminus \{0\}.$

4 Proof of Theorem 1.3

The proof of the first statement relies on Theorem 4.1 [3] and in part on some arguments of the proof of Theorem 5.1 therein. The main additional difficulty is that our reference exponential parameter θ_0 is *a priori* larger than the one considered in this reference. We overcome this difficulty in Lemma 4.1 below. Another difficulty is that a key part of the argument relies on the assumption (E1-4) in [4], which does not apply directly when the Lyapunov parameter θ_a (in Proposition 1.5) is only assumed strictly smaller than θ_0 . To overcome this difficulty, we prove in Proposition A.1 that this last condition actually entails the assumption (E1-4) in [4].

Fix $a \in (1, \varsigma)$ such that $(\lambda^*)^a < \theta_0$ and set $H = Q_a / \inf Q_a$. According to Proposition 1.5 and since $H(z)$ tends to ∞ when $|z| \to +\infty$, there exists $\theta_a < \theta_0$ such that

$$
\limsup_{|z| \to +\infty} \frac{\mathbb{E}_z(H(Z_1) \mathbf{1}_{Z_1 \neq 0})}{H(z)} \le \theta_a < \gamma := (\theta_a + \theta_0) / 2 < \theta_0. \tag{18}
$$

In particular, there exist only finitely many communication classes $E_1, \ldots, E_{k_0}, k_0 \geq 1$, for the process $(Z_n)_{n \in \mathbb{N}}$ such that

$$
\forall i \in \{1, \ldots, k_0\}, \exists z \in E_i, \text{ such that } \frac{\mathbb{E}_z(H(Z_1) \mathbf{1}_{Z_1 \neq 0})}{H(z)} > \gamma.
$$

Note that $k_0 \neq 0$ (otherwise $\theta_0 \leq \gamma$). For all $i \in \{1, ..., k_0\}$, we define $H_i = H\mathbf{1}_{E_i}$ and denote by *Y*^(*i*) the process with state space $E_i \cup \{0\}$ defined by

$$
Y_n^{(i)} = \begin{cases} Z_n & \text{if } Z_n \in E_i \\ 0 & \text{otherwise.} \end{cases}
$$

For all $i \in \{1,\ldots,k_0\}$, we denote by $\theta_{0,i}$ the absorption parameter of $Y^{(i)}$:

$$
\theta_{0,i} = \sup_{z \in E_i} \sup \{ \theta > 0, \, \liminf_{n \to +\infty} \theta^{-n} \mathbb{P}_z(Y_n^{(i)} \neq 0) > 0 \}
$$

and we set

,

$$
\bar{\theta} = \max_{i \in \{1,\dots,k_0\}} \theta_{0,i}.
$$

We assume that $E_1, \ldots, E_k, k \in \{1, \ldots, k_0\}$, are the communication classes such that $\theta_{0,i} = \bar{\theta}$, so that $\max_{i \geq k+1} \theta_{0,i} < \bar{\theta}$.

We define the set

$$
E'_0 := \mathbb{N}^p \setminus \left(\{0\} \cup \bigcup_{i=1}^{k_0} E_i \right).
$$

By definition of E_1, \ldots, E_{k_0} , we have

$$
\forall z \in E'_0, \ \frac{\mathbb{E}_z(H(Z_1)\mathbf{1}_{Z_1\neq 0})}{H(z)} \le \gamma. \tag{19}
$$

In what follows, we make use of the following crucial lemma.

Lemma 4.1. *We have* $\bar{\theta} = \theta_0$.

Proof of Lemma 4.1. Our aim is to show that, for all $\rho > \bar{\theta} \vee \gamma$ and all $z \in \mathbb{N}^p$, we have

$$
\liminf_{n \to +\infty} \rho^{-n} \mathbb{P}_z(Z_n \neq 0) = 0. \tag{20}
$$

This shows that $\bar{\theta} \vee \gamma \ge \theta_0$ and hence $\bar{\theta} \ge \theta_0$, while the converse inequality is trivial, concluding the proof of Lemma 4.1.

For all $i \neq j \in \{1, \ldots, k_0\}$, we have

$$
\forall z \in E_j, \mathbb{P}_z(\exists n \ge 0, Z_n \in E_i) > 0 \text{ or } \forall z \in E_j, \mathbb{P}_z(\exists n \ge 0, Z_n \in E_i) = 0
$$

If the condition on the left is satisfied, we write $E_i \prec E_j$, otherwise we write $E_i \nprec E_j$. This defines a partial order on $\{1,\ldots,k_0\}$. For all $i \in \{1,\ldots,k_0\}$, we denote by $c(i)$ the maximal length of sets of distinct indices $i_1 = i, \ldots, i_c$ such that $E_{i_c} \prec \cdots \prec E_{i_2} \prec E_{i_1}$. We also define, for all $\ell \ge 1$, the set

$$
E_0^{\ell} := \{x \in E'_0, \exists i \in \{1, \dots, k_0\} \text{ with } c(i) = \ell \text{ and } x \to E_i, \text{ and } x \nrightarrow E_j \forall j \text{ such that } c(j) > \ell\},\
$$

where $x \to E_i$ means that $\mathbb{P}_x(T_{E_i} < \infty) > 0$, while $x \not\to E_j$ means that $\mathbb{P}_x(T_{E_i} < \infty) = 0$. We also set

$$
E_0^0 := E_0' \setminus \cup_{\ell=1}^{\max_i c(i)} E_0^{\ell}.
$$

We obtain (20) for all $Z \in \mathbb{N}^p \setminus \{0\}$ by proving that, for any fixed $\rho > \bar{\theta} \vee \gamma$,

$$
\mathbb{E}_z(H(Z_n)\mathbf{1}_{Z_n\neq 0}) \le C_\rho \rho^n H(z), \ \forall z \in \mathbb{N}^p \setminus \{0\} \tag{21}
$$

where C_ρ does not depend on *z*. This last inequality is immediate for all $z \in E_0^0$ (see beginning of Step 1 below) and is proved for all $z \in E_i \cup E_0^{c(i)}$ $\hat{a}_0^{(i)}$, $i \in \{1, ..., k_0\}$, by induction on $c(i)$.

Step 1. We first observe that, for all $\ell \ge 0$ and all $z \in E_0^{\ell}$, using (19),

$$
\mathbb{E}_z\left(H(Z_n)\mathbf{1}_{Z_n \in E_0^{\ell}}\right) \le \gamma^n H(z) \le \rho^n H(z). \tag{22}
$$

In particular, since $E_0^0 \cup \{0\}$ is a closed set, this shows (21) for all $z \in E_0^0$.

We now prove that, for all $i \in \{1, ..., k_0\}$ and all $z \in E_i$,

$$
\mathbb{E}_z(H(Z_n)\mathbf{1}_{Z_n\in E_i}) \le C\rho^n H(z). \tag{23}
$$

In order to do so, we consider the set

$$
K_i = \{ z \in E_i, \ \frac{\mathbb{E}_z(H(Z_1)\mathbf{1}_{Z_1\neq 0})}{H(z)} > \gamma \}.
$$

For all $z \in E_i$, since $\rho > \bar{\theta}$ and by definition of $\bar{\theta}$,

$$
\liminf_{n \to +\infty} \theta^{-n} \mathbb{P}_z(Y_n^{(i)} \neq 0) = 0. \tag{24}
$$

Since K_i is finite according to (18), we deduce that there exists a constant $C > 0$ such that

$$
\mathbb{P}_z(Z_n \notin E_i) = \mathbb{P}_z(Y_n^{(i)} \neq 0) \le C\rho^n, \forall z \in K_i.
$$
 (25)

In what follows, the constant *C* may change from line to line. Moreover, denoting $\tau_{K_i} = \min\{\ell \geq 1\}$ 0, $Z_{\ell} \in K_i$, we deduce from the definition of *H* and of K_i that

$$
\mathbb{P}_z(Z_n \notin E_i \text{ and } n < \tau_{K_i}) \le \mathbb{E}_z(H(Z_n) \mathbf{1}_{Z_n \notin E_i \text{ and } n < \tau_{K_i}}) \le \gamma^n H(z), \ \forall z \in E_i. \tag{26}
$$

Hence, for all $z \in E_i$ and all $n \ge 0$, we have, using the strong Markov property at time τ_{K_i} ,

$$
\mathbb{P}_z(Z_n \notin E_i) \le \mathbb{P}_z(Z_n \notin E_i \text{ and } n < \tau_{K_i}) + \sum_{\ell=0}^n \mathbb{P}_z(\tau_{K_i} = \ell) \sup_{y \in K_i} \mathbb{P}_y(Z_{n-\ell} \notin E_i)
$$

\n
$$
\le \gamma^n H(z) + \sum_{\ell=0}^n \mathbb{P}_z(\tau_{K_i} = \ell) C \rho^{n-\ell}
$$

\n
$$
\le \gamma^n H(z) + \sum_{\ell=0}^n \mathbb{P}_z(Z_\ell \ne 0 \text{ and } \ell - 1 < \tau_{K_i}) C \rho^{n-\ell}
$$

\n
$$
\le \gamma^n H(z) + \sum_{\ell=0}^n \gamma^{\ell-1} H(z) C \rho^{n-\ell}
$$

\n
$$
\le C \rho^n H(z),
$$

where *C* may depend on ρ and γ . Finally, iterating (10) and using the last inequality, we deduce that

$$
\begin{aligned} \mathbb{E}_z(H(Z_n)1_{Z_n \notin E_i}) &\leq \theta_a^n H(z) + \sum_{\ell=0}^{n-1} \frac{C_a}{\inf Q_a} \theta_a^{n-\ell-1} \mathbb{P}_z(Z_\ell \neq 0) \\ &\leq C \gamma^n H(z) + \sum_{\ell=0}^{n-1} \frac{C_a}{\inf Q_a} \gamma^{n-\ell-1} C \rho^\ell H(z) \\ &\leq C \rho^n H(z), \end{aligned}
$$

where *C* may depend on ρ and γ . This proves (23) for all $z \in E_i$.

Step 2. Assume that there exists $c \ge 0$ such that (21) holds true for all $z \in E_0^0, \ldots, E_0^c$ and all $z \in E_i$ with $c(i) \leq c$ (Steps 1 shows that this is true for $c = 0$). If $c = \max_{i \in \{1, ..., k_0\}} c(i)$, then there is nothing to prove. Otherwise, let $i \in \{1, ..., k_0\}$ such that $c(i) = c + 1$.

Denote by T_i the first exit time from E_i by $(Z_n)_{n \in \mathbb{N}}$ and let us use a similar computation as in the end of Step 1. For all $z \in E_i$ and all $n \ge 0$, we have

$$
\mathbb{E}_z(H(Z_n)1_{Z_n\neq 0}) = \mathbb{E}_z(H(Z_n)1_{Z_n\in E_i}) + \sum_{\ell=0}^n \mathbb{E}_x\left(1_{T_i=\ell} \mathbb{E}_{X_\ell}(H(Z_{n-\ell})1_{Z_{n-\ell}\neq 0})\right)
$$

$$
\leq C\rho^n H(z) + \sum_{\ell=0}^n \mathbb{E}_x\left(1_{T_i=\ell} C_\rho \rho^{n-\ell} H(X_\ell)\right)
$$

where we used (23), the fact that

$$
Z_{T_i} \in \{0\} \cup \bigcup_{\ell \le c} E_0^{\ell} \cup \bigcup_{c(j) \le c} E_j \text{ almost surely},
$$

and the induction assumption. We then obtain

$$
\mathbb{E}_z(H(Z_n)1_{Z_n\neq 0}) \le C\rho^n H(z) + \sum_{\ell=0}^n \mathbb{E}_x \left(1_{T_i > \ell-1} C_\rho \rho^{n-\ell} (\theta_a H(X_{\ell-1}) + C_a) \right)
$$

$$
\le C\rho^n H(z) + \sum_{\ell=0}^n \mathbb{E}_x \left(1_{T_i > \ell-1} C_\rho \rho^{n-\ell} (\theta_a + C_a) H(X_{\ell-1}) \right)
$$

where we used the Markov property at time *ℓ*−1 and Proposition 1.5. We deduce from (23) that

$$
\mathbb{P}_z(Z_n \neq 0) \le C\rho^n H(z) + C_\rho(\theta_a + C_a)C \sum_{\ell=0}^n \rho^n H(z)
$$

$$
\le C(n+1)\rho^n H(z).
$$

Since this is true for any *ρ* > $\bar{\theta}$ v γ , this proves (21) for all *z* ∈ *E_i* with *c*(*i*) = *c* + 1.

The same arguments show that this also holds true for all $z \in E_0^{c+1}$.

This concludes the proof of Lemma 4.1.

From now on, we assume without loss of generality that there exists $k \in \{1, ..., k_0\}$ such that, for all $i \in \{k+1, ..., k\}$, $\theta_{0,i} < \bar{\theta}$, and, for all $i \in \{1, ..., k\}$, $\theta_{0,i} = \bar{\theta} = \theta_0$. For any $i \in \{k+1, ..., k_0\}$, we have, fixing any $\rho\in(\max_{j\in\{k'+1,...,k\}}\theta_{0,j},\bar{\theta})$ and using the same method as in Step 1 from the proof of Lemma 4.1,

$$
\mathbb{E}_z(H(Z_n)1_{n
$$

where we recall that T_i is the first exit time from E_i by $(Z_n)_{n\in\mathbb{N}}$ and where C_ρ is a positive constant which may depend on *ρ*. Setting

$$
E_0 := E'_0 \cup \bigcup_{i=k+1}^{k_0} E_i,
$$

one easily checks, using the last inequality and (19), that, for any $r \in (\gamma \vee \rho, \bar{\theta})$, there exists $C_r > 0$ such that

$$
\mathbb{E}_z(H(Z_n)1_{Z_n\in E_0})\leq C_r r^n H(z),\; \forall z\in E_0.
$$

In order to apply Theorem 4.1 in [3], which implies the first part of Theorem 1.3, it remains to prove that, for all $i \in \{1, ..., k'\}$ and all $z \in E_i$, we have

$$
\left|\theta_0^{-n}\mathbb{E}_z\left(f(Z_n)\mathbf{1}_{Z_n\in E_i}\right)-h_i(z)\mu_i(f)\right|\leq C\alpha^n H_i(x)\left\|f\right\|_{H_i},\quad\forall\,n\geq 0,\,\forall\,f\in L^\infty(H_i),\tag{27}
$$

for some $C > 0$, $\alpha \in (0,1)$, some non-negative non-zero function h_i and some probability measure μ_i on E_i . In order to do so, its is sufficient to check Assumptions (E1-4) in Section 2 of [4]. These conditions are recalled in the Appendix and we will more precisely check Assumptions $(E1)$, $(E2')$, $(E3)$ and $(E4)$, which is sufficient to obtain $(E1-4)$ according to Proposition A.1. Fixing $i \in \{1, ..., k'\}$, we define the set

$$
K_i:=\left\{z\in E_i,\; \frac{\mathbb{E}_z(H_i(Z_1)\mathbf{1}_{Z_1\neq 0})}{H_i(z)}>\gamma\right\},
$$

which is finite according to (18).

Since E_i is a communication class and the process is assumed to be aperiodic, there exists $n_1^i \geq 0$ such that

$$
c_1^i := \inf_{x,y \in K_i} \mathbb{P}_x(Z_{n_1^i} = y) > 0.
$$

Note that $K_i \neq \emptyset$, since it would imply that $\theta_{0,i} = \theta_0 \leq \gamma$. Fix $z_i \in K_i$. Then, for all $z \in K_i$,

$$
\mathbb{P}_z(Z_{n_1^i} \in \cdot) \ge \mathbb{P}_z(Z_{n_1^i} = z_i) \delta_{z_i} (\cdot) \ge c_1^i v_i(\cdot \cap K),
$$

for $v_i = \delta_{z_i}$. This entails (E1). In addition, taking $\varphi_1 = H_i$ and $\theta_1 = \gamma < \theta_0 = \theta_{0,i}$, the condition (E'2) is an immediate consequence of the definition of *Kⁱ* .

For all $x, y \in K_i$ and $n \ge 0$, we have, using the Markov property at time n_1^i and the fact that $\mathbb{P}_x(Z_n \neq 0)$ is non-increasing in *n* (as {0} is absorbing),

$$
\mathbb{P}_x(Z_n \neq 0) \ge \mathbb{P}_x(Z_{n \vee n_1^i} \neq 0) \ge \mathbb{P}_x(Z_{n_1^i} = y) \mathbb{P}_y(Z_{n \vee n_1^i - n_1^i} \neq 0) \ge c_1^i \mathbb{P}_y(Z_n \neq 0).
$$

This proves (E3).

Finally (E4) holds true since we assumed that the process *Z* is aperiodic. This concludes the proof of the first part of Theorem 1.3.

To conclude the proof of Theorem 1.3, it remains to prove that the set $E = \{x \in \mathbb{N}^p \setminus \{0\}, \sum_{i \in I} \eta_i(x) = 0\}$ 0} is finite and that any quasi-stationary distribution with absorption parameter strictly smaller than θ_0 is supported by *E*.

We first observe that Theorem 4.1 in [3] also implies that $j \equiv 0$ on the set *E*. Assume on the contrary that *E* is not finite and fix $z_0 \in \mathbb{N}^p \setminus \{0\}$ such that $\sum_{i=1}^{\ell} \eta_i(z_0) > 0$ and $j(z_0) = 0$ (existence of such a point *z*⁰ is guaranteed by Proposition 2.3 in [3]). We observe that

$$
\liminf_{n\to+\infty}\theta_0^{-n}\mathbb{P}_{z_0}(Z_n\neq 0)>0.
$$

Then, by primitivity of M , there exists n_0 such that for all $x \in N^p$ we have $M^{n_0}(x) > 0$. This implies that for $k > 0$ large enough we have $P_{kx}(Z_{n_0} \ge z_0) > 0$. Indeed, if it is not the case, we can conclude using Theorem 6 in [9] that

$$
\mathfrak{M}^{n_0}(x) = \lim_{k \to +\infty} \frac{\mathbb{E}(Z_{n_0} \mid Z_0 = kx)}{k} \ge \lim_{k \to +\infty} \frac{z_0}{k} = 0,
$$

which is a contradiction. Then, we have for any *z* large enough $\mathbb{P}_z(Z_{n_0} \ge z_0) > 0$. Hence, using the fact that *E* is not finite and hence contains arbitrarily large points, there exists $z \in E$ and $z'_0 \ge z_0$ such that $P_z(Z_{n_0} = z'_0)$ $\binom{1}{0}$ > 0. By super-additivity of *ξ*, we have $\mathbb{P}_{z_0'}(Z_n \neq 0) \ge \mathbb{P}_{z_0}(Z_n \neq 0)$ and hence

$$
\liminf_{n\to+\infty}\theta_0^{-n}\mathbb{P}_z(Z_{n_0+n}\neq 0)>0.
$$

This is not compatible with the definition of *E* and (5). We have thus proved by contradiction that *E* is finite.

Let v_{OS} be a quasi-stationary distribution for *Z* in $\mathcal{M}(Q_a)$ with absorption parameter θ_{OS} < *θ*₀. Then (5) implies that $\sum_{i=1}^{\ell}$ $\nu_{QS}(\eta_i) = 0$ and hence the support of ν_{QS} is included in *E*. Since *E* is finite and since *E* is a closed set for the Markov chain, we deduce, by usual finite dimensional linear algebra arguments, that *νQS* belongs to the convex hull of any maximal family of positive left eigenvector of the transition matrix of *Z* restricted to *E*. This concludes the proof of the penultimate assertion of Theorem 1.3.

To conclude, it remains to prove that there are no quasi-stationary distributions for *Z* in $\mathcal{M}(Q_a)$ with absorption parameter strictly larger than θ_0 . This is a direct consequence of (5). Indeed, since *j* is bounded, integrating this inequality with respect to a quasi-stationary distribution v_{OS} with absorption parameter θ_{OS} , and taking $f \equiv 1$, shows that

$$
\limsup_{n \to +\infty} \theta_0^{-n} n^{-\|j\|_{\infty}} \theta_{QS}^n \mathbb{P}_z(Z_n \neq 0) < +\infty,
$$

and hence $\theta_{QS} \leq \theta_0$.

5 Proof of Theorem 1.4

Fix $a \in (1, \varsigma)$ such that $(\lambda^*)^a < \theta_0$, and consider the Lyapunov type property with constants θ_a and C_a from Proposition 1.5. We make use of Section 2 in [4] In a first step, we check that Assumption E therein (recalled in the Appendix below) holds true for the process under consideration. In a second step, we prove (6), (7) and (8) and (9) using the results of [4].

Step 1. Assumption (E) holds true. Fix $\theta_1 \in (\theta_a, \theta_0)$. As \mathcal{P} is strictly positive on $\mathbb{R}^p_+ \setminus \{0\}$ and positively homogeneous, we can fix $r_1 \geq p$ large enough so that $Q_a(z) \geq \frac{C_a}{\theta_1 - \theta_a}$ for all $|z| \geq r_1$, so that

$$
\mathbb{E}_z(Q_a(Z_1)) \le \theta_a Q_a(z) + C_a \le \theta_1 Q_a(z), \quad \forall z \in \mathbb{N}^p \text{ such that } |z| \ge r_1.
$$
 (28)

Let us set $K = \{z \in \mathbb{N}^p \setminus \{0\}, |z| \le r_1\}$ and $\varphi_1 = Q_a / \inf Q_a$ (note that $\inf Q_a > 0$ by Assumption (P)).

We deduce from the irreducibility and aperiodicity properties that there exists $n_1 \geq 0$ such that

$$
c_1 := \inf_{x,y \in K} \mathbb{P}_x(Z_{n_1} = y) > 0. \tag{29}
$$

Setting $v = \delta_{(1,...,1)}$, this entails property (E1).

We have $\inf \varphi_1 = 1$, so that the first line of (E2) is satisfied. Moreover, according to (28), the third line of (E2) also holds true.

Choosing $\theta_2 \in (\theta_1, \theta_0)$, we deduce from the definition of θ_0 and the irreducibility of *Z* that

$$
\theta_2^{-n} \min_{i \in \{1,\dots,p\}} \mathbb{P}_{e_i}(Z_n \neq 0) \xrightarrow[n \to +\infty]{} +\infty,
$$

where e_i is the i^{th} element of the canonical basis of \mathbb{N}^p . Since, by the super-additivity of *ξ*, the process *Z* is stochastically non-decreasing in the initial condition and since, for all $z \in \mathbb{N}^p \setminus \{0\}$, there exists $i \in \{1, ..., p\}$ such that $z \geq e_i$, we deduce that

$$
\theta_2^{-n} \inf_{z \in \mathbb{N}^p \setminus \{0\}} \mathbb{P}_z(Z_n \neq 0) \ge \theta_2^{-n} \min_{i \in \{1, \dots, p\}} \mathbb{P}_{e_i}(Z_n \neq 0) \xrightarrow[n \to +\infty]{} +\infty.
$$

Let $n \ge 1$ be large enough so that θ_2^{-n} inf_{$z \in \mathbb{N}^p \setminus \{0\}$ $\mathbb{P}_z(Z_n \ne 0) \ge 1$, and set, for all $z \in \mathbb{N}^p \setminus \{0\}$}

$$
\varphi_2(z) = \frac{1}{\sum_{k=0}^{n-1} \theta_2^{-k}} \sum_{k=0}^{n-1} \theta_2^{-k} \mathbb{P}_z(Z_k \neq 0).
$$

We have, for all $z \in \mathbb{N}^p \setminus \{0\}$ and using the Markov property at time 1,

$$
\left(\sum_{k=0}^{n-1} \theta_2^{-k}\right) \mathbb{E}_z(\varphi_2(Z_1) \mathbf{1}_{Z_1 \neq 0}) = \sum_{k=0}^{n-1} \theta_2^{-k} \mathbb{P}_z(Z_{k+1} \neq 0)
$$

= $\theta_2 \sum_{k=1}^n \theta_2^{-k} \mathbb{P}_z(Z_k \neq 0)$
= $\theta_2 \left(\sum_{k=0}^{n-1} \theta_2^{-k}\right) \varphi_2(z) + \theta_2 \left(\theta_2^{-n} \mathbb{P}_z(Z_n \neq 0) - 1\right)$
 $\geq \theta_2 \left(\sum_{k=0}^{n-1} \theta_2^{-k}\right) \varphi_2(z).$

As a consequence the second and fourth lines of (E2) are satisfied. Note also that we have $\inf_{\mathbb{N}^p \setminus \{0\}} \varphi_2 \geq \frac{1}{\sum_{k=0}^{n-1} \theta_2^{-k}}.$

For all *x*, *y* ∈ *K* and *n* ≥ 0, we have, using the Markov property at time *n*₁ and the fact that $\mathbb{P}_x(Z_n \neq 0)$ is non-increasing in *n* (as {0} is absorbing),

$$
\mathbb{P}_x(Z_n \neq 0) \ge \mathbb{P}_x(Z_{n \vee n_1} \neq 0) \ge \mathbb{P}_x(Z_{n_1} = y) \mathbb{P}_y(Z_{n \vee n_1 - n_1} \neq 0) \ge c_1 \mathbb{P}_y(Z_n \neq 0).
$$

This proves (E3).

Finally (E4) holds true since we assumed that the process *Z* is aperiodic and irreducible.

We have thus proved that Assumption (E) holds true for Z absorbed when it leaves $\mathbb{N}^p \setminus \{0\}$. This concludes the first step.

Step 2. Conclusion of the proof. Theorem 2.1 in [4] implies that there exists a unique quasistationary distribution v_{OS} in $\mathcal{M}(Q_a)$ and that (8) holds true (using the fact that in our case, φ_2 is lower bounded away from 0). Theorem 2.3 in this reference also implies that there exists an associated non-negative eigenfunction $\eta_{QS} \in L^\infty(Q_a)$ with eigenvalue $\hat\theta_{QS}>0$, where θ_{QS} is the absorption parameter associated to v_{OS} , such that $\inf_K \eta_{OS} > 0$ and

$$
\theta_{QS}^{-n} \mathbb{P} \cdot (Z_n \neq 0) \xrightarrow[n \to +\infty]{L^{\infty}(Q_a)} \eta_{QS}(\cdot).
$$
\n(30)

Since the process is irreducible, for all $z \in \mathbb{N}^p \setminus \{0\}$, there exists $n \ge 0$ such that $\mathbb{P}_z(Z_n = (1, \ldots, 1))$ 0 and hence

$$
\theta_{QS}^n \eta_{QS}(z) = \mathbb{E}_z(\eta_{QS}(Z_n) \mathbf{1}_{Z_n \neq 0}) \ge \eta_{QS}((1,\ldots,1)) \mathbb{P}_z(Z_n = (1,\ldots,1)) > 0,
$$

so that η_{QS} is positive. We thus proved (6) (for the absorption parameter θ_{QS} which we prove below to be equal to θ_0). Finally, since ξ is super-additive, $\theta_{QS}^{-n}P_z(Z_n \neq 0)$ and hence η_{QS} increases with *z*, so that η_{OS} is lower bounded away from 0. Corollary 2.7 in [4] shows that (7) holds true, and, taking $f \equiv 1$, also implies that $\theta_{OS} = \theta_0$. Finally, (9) is an immediate consequence of Corollaire 2.11 in [4]. This concludes the proof of Theorem 1.4.

A Sufficient criterion for the exponential convergence to a quasi-stationary distribution

In Section 2 of [4], the authors state that Assumption (E) is sufficient to prove exponential convergence toward a quasi-stationary distribution. Let us recall this assumption, with the notations and settings of the present paper.

In what follows, we consider a process $(Z_n)_{n\in\mathbb{N}}$ evolving in a measurable state space $E \cup \{\partial\}$, where *∂* ∉ *E* is an absorbing point.

Assumption (E). There exist a positive integer n_1 , positive real constants $\theta_1, \theta_2, c_1, c_2, c_3$, two functions $\varphi_1, \varphi_2 : E \to \mathbb{R}_+$ and a probability measure *v* on a measurable subset $K \subset E$ such that

(E1) *(Local Dobrushin coefficient).* ∀*z* ∈ *K*,

$$
\mathbb{P}_z(Z_{n_1}\in \cdot)\geq c_1\nu(\cdot\cap K).
$$

(E2) *(Global Lyapunov criterion)*. We have $\theta_1 < \theta_2$ and

$$
\inf_{z \in E} \varphi_1(z) \ge 1, \sup_{z \in K} \varphi_1(z) < \infty
$$
\n
$$
\inf_{z \in K} \varphi_2(z) > 0, \sup_{z \in E} \varphi_2(z) \le 1,
$$
\n
$$
\mathbb{E}_z(\varphi_1(Z_1) \mathbf{1}_{Z_1 \neq \partial}) \le \theta_1 \varphi_1(z) + c_2 \mathbf{1}_K(z), \forall z \in E
$$
\n
$$
\mathbb{E}_z(\varphi_2(Z_1) \mathbf{1}_{Z_1 \neq \partial}) \ge \theta_2 \varphi_2(z), \forall z \in E.
$$

(E3) *(Local Harnack inequality).* We have

$$
\sup_{n\in\mathbb{Z}_+}\frac{\sup_{y\in K}\mathbb{P}_y(Z_n\neq \partial)}{\inf_{y\in K}\mathbb{P}_y(Z_n\neq \partial)}\leq c_3.
$$

(E4) *(Aperiodicity).* For all $z \in K$, there exists $n_4(z)$ such that, for all $n \ge n_4(z)$,

$$
\mathbb{P}_z(Z_n \in K) > 0.
$$

Our aim is to show that the condition involving the function φ_2 can be replaced by a condition involving θ_0 , where

$$
\theta_0 = \sup_{z \in E} \sup \{ \theta > 0, \liminf_{n \to +\infty} \theta^{-n} \mathbb{P}_z(Z_n \neq \partial) > 0 \},\tag{31}
$$

More precisely, we consider the following assumption.

(E2') We have $\theta_1 < \theta_0$ and

$$
\inf_{z \in E} \varphi_1(z) \ge 1, \sup_{z \in K} \varphi_1(z) < \infty
$$
\n
$$
\mathbb{E}_z(\varphi_1(Z_1) \mathbf{1}_{Z_1 \neq \partial}) \le \theta_1 \varphi_1(z) + c_2 \mathbf{1}_K(z), \forall z \in E.
$$
\n(32)

Proposition A.1. *Assume that there exist a positive integer* n_1 *, positive real constants* θ_1 , c_1 , c_2 , c_3 , a function $\varphi_1 : E \to \mathbb{R}_+$ *and a probability measure* v *on a measurable subset* $K \subset E$ *such that* (E1), *(E'2), (E3) and (E4) hold true. Then, for any* $\theta_2 \in (\theta_1, \theta_0)$ *, there exists* φ_2 *such that Assumption (E2) holds true.*

Proof. In a first step, we prove that, for any $\theta_2 \in (\theta_1, \theta_0)$, there exists $n \ge 1$ such that $\inf_{z \in K} \theta_2^{-n} \mathbb{P}_z(Z_n \in K)$ *K*) \geq 1. In a second step, we conclude by building a function φ_2 which satisfies Assumption (E2).

Step 1. Assume on the contrary that there exists $\theta_2 \in (\theta_1, \theta_0)$ such that, for all $n \ge 1$, $\inf_{z \in K} \theta_2^{-n} \mathbb{P}_z(Z_n \in \theta_1)$ *K*) < 1. Then, using (E1), we deduce that, for all $n \ge n_1$,

$$
1 > \inf_{z \in K} \theta_2^{-n} \mathbb{P}_z(Z_n \in K) \ge \theta_2^{-n} c_1 \mathbb{P}_v(Z_{n-n_1} \in K)
$$

and hence that, for all $n \geq 0$,

$$
\mathbb{P}_{\nu}(Z_n \in K) < \frac{\theta_2^{n_1}}{c_1} \theta_2^n.
$$

Applying (32) iteratively, we deduce that, for all $z \in E$, for all $n \ge 1$,

$$
\mathbb{E}_z(\varphi_1(Z_n)\mathbf{1}_{Z_n\neq\partial})\leq\theta_1^n\varphi_1(z)+c_2\sum_{k=1}^n\theta_1^{n-k}\mathbb{P}_z(Z_{k-1}\in K).
$$

Integrating with respect to *ν* and using the two previous inequalities, we get

$$
\mathbb{E}_{\nu}(\varphi_1(Z_n)\mathbf{1}_{Z_n\neq\partial}) \leq \theta_1^n \nu(\varphi_1) + c_2 \sum_{k=1}^n \theta_1^{n-k} \mathbb{P}_{\nu}(Z_{k-1} \in K) \leq \theta_1^n \nu(\varphi_1) + c_2 \sum_{k=1}^n \theta_1^{n-k} \frac{\theta_2^{n-1}}{c_1} \theta_2^k \leq C' \theta_2^n,
$$
\n(33)

for some constant *C'* where we used the fact that $v(\varphi_1) \le \sup_K \varphi_1 < +\infty$ and $\theta_1 < \theta_2$.

For all $z \in E$ and all $n \ge 0$, we have, denoting $\tau_K = \min\{k \ge 0, Z_k \in K\}$ and using the strong Markov property at time *τ^K* ,

$$
\mathbb{P}_z(Z_n \neq \partial) \leq \mathbb{E}_z(\varphi_1(Z_n)1_{Z_n \neq \partial} 1_{n < \tau_K}) + \sum_{k=0}^n \mathbb{P}_z(\tau_K = k) \sup_{y \in K} \mathbb{P}_y(Z_{n-k} \neq \partial)
$$
\n
$$
\leq \theta_1^n \varphi_1(z) + \sum_{k=0}^n \mathbb{P}_z(\tau_K = k) c_3 \mathbb{P}_v(Z_{n-k} \neq \partial),
$$

where we used (32) for the first term in the right hand side, and (E3) for the second term. Since we also have, by (32) and for all $k \ge 1$ and trivially for $k = 1$,

$$
\mathbb{P}_z(\tau_K = k) \le \mathbb{P}_z(Z_{k-1} \neq \hat{\sigma}, \ k-1 < \tau_K) \le \mathbb{E}_z(\varphi_1(Z_{k-1}) 1_{Z_{k-1} \neq \hat{\sigma}} 1_{k-1 < \tau_K}) \le \theta_1^{k-1} \varphi_1(z),
$$

and since $\mathbb{P}_{\nu}(Z_{n-k} \neq \partial) \leq \mathbb{E}_{\nu}(\varphi_1(Z_{n-k})\mathbf{1}_{Z_{n-k}\neq \partial})$, we deduce from the previous inequality and from (33) that

$$
\mathbb{P}_z(Z_n \neq \partial) \le \left(\theta_1^n + \sum_{k=0}^n \theta_1^{k-1} c_3 C' \theta_2^{n-k}\right) \varphi_1(z).
$$

In particular, for all $z \in E$, for all $\theta > \theta_2 > \theta_1$,

$$
\theta^{-n} \mathbb{P}_z(Z_n \neq \partial) \xrightarrow[n \to +\infty]{} 0,
$$

then, by definition of θ_0 , we get $\theta_0 \le \theta_2$, which contradicts assumption $\theta_2 \in (\theta_1, \theta_0)$.

This concludes the proof of our first step: for any $\theta_2 \in (\theta_1, \theta_0)$, there exists $n \ge 1$ such that $\inf_{z \in K} \theta_2^{-n} \mathbb{P}_z(Z_n \in K) \geq 1.$

Step 2. Fix $\theta_2 \in (\theta_1, \theta_0)$ and let $n \ge 1$ such that $\inf_{z \in K} \theta_2^{-n} \mathbb{P}_z(Z_n \in K) \ge 1$, and define the function

$$
\varphi_2: z\in E\mapsto C_{\theta_2}\sum_{k=0}^{n-1}\theta_2^{-k}\mathbb P_z(Z_k\in K)
$$

with $C_{\theta_2} := \left(\frac{1-\theta_2^n}{1-\theta_2}\right)$ \int_0^{-1} . Then, $\inf_{z \in K} \varphi_2(z) > 0$ and $\sup_{z \in E} \varphi_2(z) \le 1$. In addition,

$$
\mathbb{E}_z(\varphi_2(Z_1)\mathbf{1}_{Z_1\neq\partial}) = C_{\theta_2} \sum_{k=0}^{n-1} \theta_2^{-k} \mathbb{P}_z(Z_{k+1} \in K) = C_{\theta_2} \theta_2 \sum_{k=1}^n \theta_2^{-k} \mathbb{P}_z(Z_k \in K)
$$

= $\theta_2 \varphi_2(z) + C_{\theta_2} \theta_2 \left(\theta_2^{-n} \mathbb{P}_z(Z_n \in K) - \mathbf{1}_K(z)\right).$

Since $\theta_2^{-n} \mathbb{P}_z(Z_n \in K) - \mathbf{1}_K(z) \ge 0$ for all $z \in K$ (by definition of *n*) and for all $z \in E \setminus K$ (since for such *z*, we have $1_K(z) = 0$, we deduce that

$$
\mathbb{E}_z(\varphi_2(Z_1)\mathbf{1}_{Z_1\neq\partial})\geq \theta_2\varphi_2(z),\ \forall\,z\in E.
$$

This concludes the proof of Proposition A.1.

 \Box

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