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Compositional discriminant analysis through calibrated evidence functions

Paul-Gauthier Noé, Andreas Nautsch, Driss Matrouf, Pierre-Michel Bousquet and Jean-François Bonastre

Overview

Compositional discriminant analysis is a non-linear discriminant function where the discriminant components form a calibrated isometric-log-ratio transformed likelihood function over the set of classes. The latter can be called *evidence function* and can be seen as a multiclass extension of the weight-of-evidence in Bayesian updating. The compositional nature of the approach is on the treatment of the likelihood function (not on the modelled data) allowing a novel analysis method.

Considering a set of classes $\mathcal{C} = \{c_i\}_{1 \leq i \leq N}$ *and an individual who wants to infer the class of an observation* x*, its posterior probabilities are given by:*

> $P(x) = w(x) \oplus \pi$, \iff ilr $P(x) =$ ilr $w(x)$ | {z } *Evidence function* $+$ ilr π ,

where $\bm{P}(x) = \left[P(c_i \mid x) \right]_{1 \leq i \leq N'} \bm{w}(x) = \left[P(x \mid c_i) \right]_{1 \leq i \leq N}$ and $\bm{\pi} = \left[P(c_i) \right]_{1 \leq i \leq N}$ *are respectively the posterior distribution, the likelihood function and the prior.*

Bayesian updating on the Aitchison simplex [1]

where \boldsymbol{e}_i is the i^{th} vector of the standard canonical basis of \mathbb{R}^{N-1} *.*

The idempotence and the distribution of evidence function

An evidence function $\bm{l}(x) = \text{ilr}\,\bm{w}(x)$ *is calibrated if it results in the same posterior probabilities whether* $\mathbf{l}(x)$ *or the data* x *is given* [2]:

 $\forall i \in [1, N], \ P(c_i \mid x) = P(c_i \mid \boldsymbol{l}(x)) \iff \boldsymbol{l}(\boldsymbol{l}) = \boldsymbol{l}(x)$ (idempotence),

leading to the following proposition extending a well-known result for LLRs [3],

Find a diffeomorphism $g^{-1} : \mathcal{X} \mapsto \mathcal{Z}$, such that

Proposition 1 [2]: If
$$
l \mid c_1 \sim \mathcal{N}(\mu_1, \Sigma)
$$
, then $\forall i \in [2, N]$, $l \mid c_i \sim \mathcal{N}(\mu_i(\Sigma), \Sigma)$,
where $\mu_1 = \mu_1(\Sigma) = A^{-1}B \text{ vec}(\Sigma)$, and $\mu_i(\Sigma) = \mu_1(\Sigma) - \Sigma a_{i-1} - \sum_{j=1}^{i-2} \frac{1}{j+1} \Sigma a_j$,
 $A \in \mathcal{M}_{N-1,N-1}(\mathbb{R})$ and $B \in \mathcal{M}_{N-1,(N-1)^2}(\mathbb{R})$ are constant matrices, and $a_i = \sqrt{\frac{i+1}{i}}e_i$

$\bm{C}\left(\bm{\Sigma}\right) = 0$ $\mathbf{0}_{d-N+1,N-1}$ \boldsymbol{I}_{d-N+1} ,

Normalizing flow—a cascade of invertible neural-networks [5]—is used to learn the mapping g through data likelihood maximization:

where the mapping's parameters θ_q and the log-Cholesky parametrized Σ are learned with automatic differentiation and gradient descent.

 -2 0
First CDA's residual co -7.5 -5.0 -2.5 0.0 2.5 5.0 7.5 -2 -1 0 1
First LDA's residual companer *Figure: Training set. Figure: Testing set in the CDA's (top) and LDA's (bottom) base space.*

Figure: Gaussian conditional densities of the likelihood function in a 3 classes case. The covariance matrix can be expressed in terms of the three divergences between each density. The densities on the ILR space (top) are with respect to the Lebesgue measure while the densities on the simplex (bottom) are with respect to the Aitchison measure [4].

The method

Notations and objective

Let,

 $\boldsymbol{x} \in \mathcal{X} \subset \mathbb{R}^d$ be an observation, $\boldsymbol{l}(\boldsymbol{x})\in\mathcal{L}\subset\mathbb{R}^{N-1}$ the evidence function, $\bm{r}(\bm{x}) \in \mathcal{R} \subset \mathbb{R}^{d-N+1}$ the residual, $\mathcal{Z} = \mathcal{L} \oplus \mathcal{R}$ the base space,

$$
\begin{aligned} \mathbf{z} &= g^{-1}(\mathbf{x}) \\ \mathbf{x} &= [z] \\ \mathbf{z} &= \begin{bmatrix} l_1, \dots, l_{N-1}, r_1, \dots r_{d-N+1} \\ \text{discriminant} \\ \text{components} \end{bmatrix} \in \mathcal{Z} \\ & \begin{aligned} & \mathbf{discriminant} \\ & \mathbf{residual} \\ & \mathbf{evidence function} \end{aligned}
$$

Class-conditional densities in the base space

- [1] Juan José Egozcue and Vera Pawlowsky-Glahn. Evidence information in bayesian updating, 2011.
- [2] Paul-Gauthier Noé.

We choose the densities in the base space according to Proposition 1:

 $\forall i \in \llbracket 1, N \rrbracket, \boldsymbol{z} \mid c_i \sim \mathcal{N}(\boldsymbol{m_i}(\boldsymbol{\Sigma}), \boldsymbol{C}(\boldsymbol{\Sigma}))$,

where the covariance matrix Σ is the **only parameter** and:

- $m_i(\Sigma)$ is the concatenation of $\mu_i(\Sigma)$ and the $(d-N+1)$ -dimensional zero vector
- the covariance matrix $C(\Sigma)$ is the following block matrix:

$$
\boldsymbol{C}(\boldsymbol{\nabla}) = \begin{bmatrix} \boldsymbol{\Sigma} & \mathbf{0}_{N-1,d-N+1} \end{bmatrix}
$$

In this way,

- the first N − 1 dimensions of z form the **evidence function**,
- the others form the **residual** normally distributed with a zero mean vector and an identity covariance matrix regardless of the class.

Learning the discriminant function with normalizing flow

$$
\log f\left(\mathcal{D}\mid\theta_g,\boldsymbol{\Sigma}\right)=\sum_{i=1}^N\left(\sum_{(x,c)\in\mathcal{D}|c=c_i}\log\left(f_{\mathcal{Z}|c_i,\boldsymbol{\Sigma}}\left(\boldsymbol{z}\right)\left|\det\left(\frac{\partial\boldsymbol{x}}{\partial\boldsymbol{z}}\right)\right|^{-1}\right)\right),
$$

Change of variable

Toy experiments

With Gaussians

Table: C*llr measures [6]. Samples from the non-concerned class are discarded.*

With MNIST

$0 / 23456789$

The flattened images reduced to 40 principal components are used as inputs

Table: Cross-entropy and accuracy measures on the testing set for the MNIST's digit recognition task.

Figure: UMAP visualization of the evidence function. Figure: UMAP visualization of the residual.

Interpolation between digits in the space of evidence functions

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Discussion

- The **discriminant components** form a calibrated isometric-log-ratio transformed likelihood function: an **evidence function**;
- No explicit assumption is made on the distribution of the data (except the existence of a diffeomorphism that would transform the data into the target Gaussians);
- **Each discriminant component expresses a weight-of-evidence** in favour of a group of classes against another group according to the used Aitchison basis;
- This can be used for uncertainty-aware decisions avoiding under or overconfident predictions;
- In addition, the discriminant space has the **intuitive** Euclidean vector space structure given by the **Aitchison geometry**, allowing interpretability, distance measures, interpolations, and straightforward computation of the posterior probability distribution by simply shifting the evidence function by the prior.

References

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- [5] Laurent Dinh, Jascha Sohl-Dickstein, and Samy Bengio. Density estimation using real nvp, 2017.
- [6] Niko Brümmer and Johan du Preez. Application-independent evaluation of speaker detection, 2006.