



HAL
open science

Experimental Comparison of Ensemble Methods and Time-to-Event Analysis Models Through Integrated Brier Score and Concordance Index

Camila Fernandez, Chung Shue Chen, Pierre Gaillard, Alonso Silva

► **To cite this version:**

Camila Fernandez, Chung Shue Chen, Pierre Gaillard, Alonso Silva. Experimental Comparison of Ensemble Methods and Time-to-Event Analysis Models Through Integrated Brier Score and Concordance Index. 2024. hal-04500089

HAL Id: hal-04500089

<https://hal.science/hal-04500089>

Preprint submitted on 11 Mar 2024

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.



Distributed under a Creative Commons Attribution 4.0 International License

Experimental Comparison of Ensemble Methods and Time-to-Event Analysis Models Through Integrated Brier Score and Concordance Index

Camila Fernandez Chung Shue Chen Pierre Gaillard Alonso Silva

Abstract

Time-to-event analysis is a branch of statistics that has increased in popularity during the last decades due to its many application fields, such as predictive maintenance, customer churn prediction and population lifetime estimation. In this paper, we review and compare the performance of several prediction models for time-to-event analysis. These consist of semi-parametric and parametric statistical models, in addition to machine learning approaches. Our study is carried out on three datasets and evaluated in two different scores (the integrated Brier score and concordance index). Moreover, we show how ensemble methods, which surprisingly have not yet been much studied in time-to-event analysis, can improve the prediction accuracy and enhance the robustness of the prediction performance. We conclude the analysis with a simulation experiment in which we evaluate the factors influencing the performance ranking of the methods using both scores.

Keywords: Ensemble methods, time-to-event analysis, integrated Brier score, concordance index.

1 INTRODUCTION

Time-to-event analysis is popular in medical research for predicting the lifetime of populations. It is also widely used in many fields in order to predict the time until a certain critical event occurs, which may be the recurrence of a disease, the customer churn in business management and operation research, recidivism in social science and psychology, the failure of machines in industrial engineering, etc. One of the most important characteristics of time-to-event analysis, which makes a significant difference from classical regression problems [50], [3], is a phenomenon known as censorship, and specifically, in this paper we treat the problem of right censorship. Right censorship arises from the fact that a study may finish before all the samples reach the critical event or because some of the individuals have withdrawn from the study before it ends. As a result, not all the samples may have reached their failure time during the observed period, such that there will be a subset of them whose observed time will represent a lower bound for the critical time.

Many different models have been proposed in order to predict survival times. One of the most widely used ones was proposed by Cox [11] in 1972; this is a semi-parametric model which is composed of an unknown baseline hazard function that depends on the time and the effect of the covariates given by an exponential function. Later, other parametric techniques were proposed by Aalen [1] and Weibull [47]. Recently, machine learning methods have attracted much attention and many non-parametric model-based machine learning techniques for time-to-event analysis have been proposed, such as random survival forest [27], DeepSurv [29], gradient boosting Cox [36] and survival support vector machine [35]. In this paper, we present a comparison of several of these models through two different scores, the concordance index [24] and the integrated Brier score [18], and among different types of data sets with the objective to study how the different models behave and compare their effectiveness.

Ensemble methods are learning algorithms that combine different models by optimizing certain weighting procedures in order to obtain a predictor that will be the combination of multiple learners. One of the main advantages of ensemble methods is the fact that they can inherit the good properties of each of the predictors and use them whenever it is most suitable, for example, if we have a dataset that behaves

better for a particular type of models, then the weighting procedure will privilege this type of models and thus leads to an increment of accuracy that is independent of the chosen dataset. Note that this can be extended to time-varying weighting by which we can also take advantage of the fact that there are some models that vary their performance over time or over the distribution (see [4]), where we ponder differently the methods that are better for predicting distribution tails and the ones for predicting the center of the distributions. Ensemble methods are well known and used in many applications of data analytics and machine learning, but it is a technique that has not been yet widely explored in time-to-event analysis (see some examples in [48]).

The existing literature lacks clean performance comparison between time-to-event analysis methods and how to calibrate parameters. Van Wieringen et al. [45] reviewed the performance of different methods applied to the particular case of gene expression data. The methods that are able to handle this type of problem, where the number of features exceeds by far the number of samples, are very specific and do not necessarily represent the general case of survival analysis problems.

Contributions. The main contribution of this paper is to give a detailed comparison of different and diverse time-to-event analysis methods using two widely used scores. The above gives us a detailed comparative study of the time-to-event analysis models and their different advantages and disadvantages. To this end, we compare the performance using three datasets and we study the impact of optimizing the hyperparameters through a randomized search. We observe that the method ranking varies across each dataset, making it challenging to select the most appropriate model without prior knowledge. To address this issue, we propose combining these different methods to enhance robustness across datasets. This is carried out by optimizing the parameters of a convex combination of the methods described in Section 2.1, such that the integrated Brier score is minimized. Finally, we conduct simulation experiments aimed at deepening insights from the dataset comparison and studying the factors influencing method performance ranking. We generate data using three different techniques under three scenarios: increasing the number of samples, reducing the number of features, and augmenting the percentage of censorship.

Paper outline. First, we present the preliminaries and definitions for our study, together with the implemented methods: Cox proportional hazard, Gradient boosting Cox, Random survival forest, Weibull accelerated failure time, Aalen’s additive and DeepSurv. In Section 3, we exhibit our implementation of ensemble methods. In Section 4, we present the three datasets (Primary biliary cirrhosis, German breast cancer and Telecom churn) used for our study. Section 5 shows the comparison of the various techniques and their numerical results. Besides, we show the performance of the ensemble method. In Section 6, we present the simulation experiment and finally, Section 7 contains some concluding remarks.

2 PRELIMINARIES

The main objective of time-to-event analysis is to estimate the distribution of survival times. Given a set of N subjects with its respective vector of covariates of dimension d , $x_i = \{x_i^1, \dots, x_i^d\} \in \mathcal{X}$, $i \in \{1, \dots, N\}$, we assume that x_i follows the distribution of a random variable X_i . Let T_i and C_i be a non-negative random variable denoting the survival and censored time, respectively. Then, we define the observed time as $Y_i = \min\{T_i, C_i\}$ and we will write $\Delta_i = \mathbb{1}\{T_i \leq C_i\}$ for the survival indicator. Under these conditions, a subject of the dataset will be described by $(x_i, y_i, \delta_i) \in \mathcal{X} \times \mathbb{R} \times \{0, 1\}$ assumed to be a realization of the random variable (X_i, Y_i, Δ_i) . In addition, we define the set of individuals at risk as $\mathcal{R}(t) = \{i \in \{1, \dots, N\} : y_i > t\}$. Let us remark that we consider that all the individuals are present at time $t = 0$. Then, the probability to survive at time t for subject i of the dataset is given by:

$$S(t|x_i) = \mathbb{P}(T_i > t | X_i = x_i).$$

In order to estimate the survival probability, many parametric and semi-parametric models assume a particular shape of the hazard function, which is defined for all $t > 0$ as:

$$\begin{aligned} h(t|x_i) &= - \frac{\partial}{\partial t} \log(S(t|x_i)) \\ &= \lim_{dt \rightarrow 0} \frac{\mathbb{P}(t \leq T_i < t + dt | T_i \geq t, X_i = x_i)}{dt}. \end{aligned}$$

We can retrieve the survival probability function by integrating the exponential of

the hazard function

$$S(t|x_i) = \exp\left(-\int_0^t h(u|x_i)du\right).$$

Each model will give us an estimator \hat{S} of the survival probability S . In addition, we define the mortality risk of an individual by a function $R : \mathcal{X} \rightarrow \mathbb{R}_+$, which will be used later to compute the concordance index. The mortality risk must satisfy $R(x_i) > R(x_j)$ if $\mathbb{P}(T_i < T_j) > 1/2$, i.e. if individual i has a higher mortality risk than individual j . Note that R is not uniquely defined and only the ranking matters. Each model will define and estimate (by providing a function \hat{R}) the mortality risk differently and we give the details separately in Section B. In addition, to measure the goodness of fit of each model, we consider two scores. Concordance Index [24] is a rank score that measures the ability of the model to correctly provide a reliable ranking of the survival times. And secondly, the integrated Brier score [18], which measures the calibration of the models by averaging the square distances between the observed survival status and the predicted survival probability. We give more details about both scores in Section A.

2.1 Methods and their implementation

We consider six methods in our study. These are Cox proportional hazard [11], gradient boosting Cox [36], random survival forest [27], Weibull AFT [47], Aalen additive [1] and DeepSurv [29]. There exist many other methods for survival analysis, such as life tables [12], different versions of cox regressions [5], [25], linear regressions [43], Bayesian network classifier based methods [17] and support vector machine [30], see [46] for more details. Nevertheless, we choose the six methods mentioned above because they are the most popular and widely used techniques, they include parametric, semi-parametric and machine learning approaches, and on the other hand, the diversity of their structure is very relevant and has a key role in ensemble methods. Note that in our implementation and the comparative study, we adopted the methods from the standard libraries: Scikit-survival [34], Lifelines [13] and PySurvival [15]. More details about the methods can be found in Section B.

3 ENSEMBLE METHODS

The main objective of ensemble methods is to combine the predictions of multiple estimators in order to improve generalizability and robustness and to obtain more reliable and accurate predictions. One has to derive effective combination rules or design powerful algorithms to boost performance. Ensemble methods consist of both empirical [22] and theoretical [39] approaches. It can be proved that weak learners can be boosted into strong learners through ensemble methods by combining multiple estimators. Applications of ensemble methods [49] can be found in many fields, such as computer vision, computer security, aided medical diagnosis, credit card fraud detection, weather forecasting, predictive maintenance, etc.

There are three reasons why it is possible to construct very good ensemble methods [14]. First, from a statistical point of view, a learner algorithm can be seen as a procedure to identify the best hypothesis space \mathcal{H} . When there is a small amount of data available, the algorithm may find many spaces that fit with the same accuracy. By aggregation, ensemble methods, however, can reduce the risk of choosing the wrong learner. Secondly, ensemble methods have computational advantages because learning algorithms can get stuck in local optimum solutions, and even when there is enough training data, it can still be challenging to find the best hypothesis. This issue can be addressed by running multiple learners from different starting points. Thirdly, in most applications of machine learning, the truth cannot be represented by any of the hypotheses in the \mathcal{H} space. However, by forming a weighted version of the elements of \mathcal{H} , it is possible to expand the space of representable functions.

In this paper, we use a gradient descent optimization algorithm to set the parameters of the convex combination of the six methods described in Section 2.1. Assuming that we have K procedures to estimate the survival probability function, let us set \hat{S}_k as the estimator proposed by the k -th method. We want to find the parameters $\lambda_k \geq 0$ such that

$$\hat{S}(t|X) = \sum_{k=1}^K \lambda_k \hat{S}_k(t|X),$$

minimizes the integrated Brier score provided that $\sum \lambda_k = 1$. In order to do this, we optimize the weights λ_k in a subset of the data \mathcal{D} of size n . We consider the

gradient vector as the descent direction, which follows the definition of integrated Brier score (see Section A.2), the j -partial derivative is given by:

$$\begin{aligned} \frac{\partial IBS(\hat{S}, \mathcal{D})}{\partial \lambda_j} &= \frac{1}{\tau n} \sum_{i=1}^n \int_0^{\tau} W_i(t) \cdot 2(\mathbb{1}\{y_i > t\} \\ &\quad - \sum_{k=1}^K \lambda_k \hat{S}_k(t|x_i)) \cdot (-S_j(t|x_i)) dt. \end{aligned} \quad (1)$$

The gradient descent algorithm is presented in Algorithm 1.

Algorithm 1 Exponential Gradient Descent

Require: T number of iteration, $\eta > 0$ learning rate

- 1: **Initialization:** $\lambda(0) = (1/K, \dots, 1/K)$
 - 2: **for** each iteration $t = 1, \dots, T$ **do**
 - 3: Define $Z_t = \sum_{k=1}^K \lambda_k(t) \exp(-\eta Df_k)$,
 - 4: where $Df_k = \frac{\partial IBS(\hat{S}, \mathcal{D})}{\partial \lambda_k}$ defined in (1).
 - 5: Update $\lambda_k(t+1) = \frac{\lambda_k(t) \exp(-\eta Df_k)}{Z_t}$ for all $k = 1, \dots, K$.
 - 6: **end for**
-

Here, we consider η a constant learning rate with initial λ equitably distributed. The iteration process is repeated until it reaches a maximum number that is set as 10000. We estimate the optimal aggregation weights each time when we fit the methods in a cross-validation process of five folds. It is important to mention that using a gradient descent algorithm for optimizing the parameters is possible thanks to the fact that the integrated Brier score function is convex, which is not the case for the concordance index.

4 DATASETS

We study three different datasets, whose general properties are summarized in Table 1.

4.1 Primary Biliary Cirrhosis (PBC)

Mayo Clinic Primary Biliary Cirrhosis dataset was made available by Therneau and Grambsch [42] and it is for studying the effects of the drug D-penicillamine on the lifetime of patients with PBC. This dataset has 276 samples and 17 covariates such

Table 1: Dataset dimensions

	Samples	Features	Censored	Percentage
PBC [42]	276	17	165	59.8 %
GBCSG2 [40]	686	8	387	56.4 %
TLCM [26]	7043	19	5174	73 %

as age, presence of ascites, cholesterol, etc. There are 165 patients who did not die at the end of the study (59.8%) and that corresponds to censored data.

4.2 German Breast Cancer Study Group 2 (GBCSG2)

German Breast Cancer Study Group was made available by Schumacher et al. [40] and it is used for studying the effects of hormone treatment on breast cancer recurrence. The dataset has 686 samples and 8 covariates, such as age, hormonal therapy, menopausal status, etc. There are 387 patients who did not get cancer again (56.4%), corresponding to censored data.

4.3 Kaggle Telco Churn (TLCM)

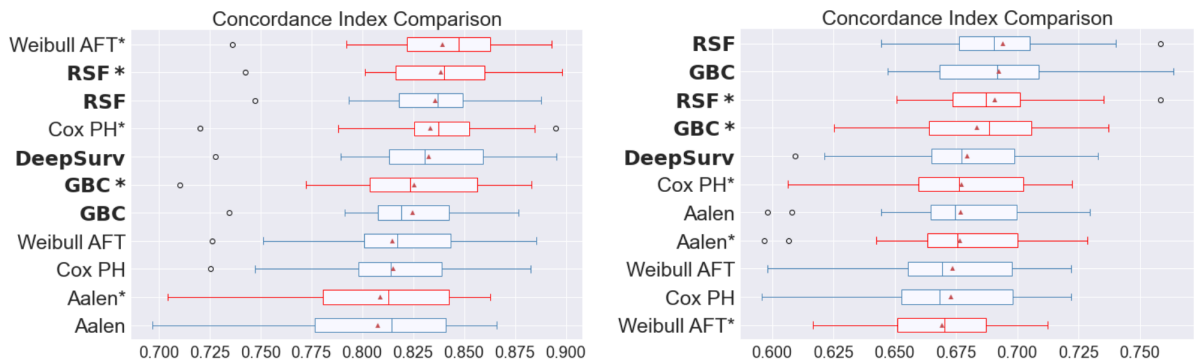
Kaggle Telco Churn dataset was made available in 2008 by Kaggle and it is a sample dataset from IBM [26]. It is used for studying the different causes of customer churn in a fictional telecommunication enterprise. The dataset has 7043 samples and 19 features such as gender, partner, dependents, phone service, etc. This dataset has 5174 clients who have not churned at the end of the study (73%) and that corresponds to censored data.

5 COMPARISON RESULTS

In the following section, we compare the six methods described in Section 2.1 through concordance index and integrated Brier score, respectively. Besides, we compare their results with that of the deployed ensemble method. For each dataset, the scores were computed 25 different times corresponding to 25 partitions (training/validation) of the dataset. This number was chosen arbitrarily in order to maintain a reasonable number of iterations without making the process too computationally expensive. Results are shown by the box plots below. Note that among Figure 1a to 4, there are some methods with their names marked with an asterisk

and their boxes colored by red, which is to indicate the implementation of a randomized search of the parameters conducted by a cross-validation process, whereas the unmarked (and blue) corresponds to adjust the method with the default parameters described in Section 2.1. In addition, the machine learning techniques were bolded to differentiate them from the semi-parametric and parametric methods.

5.1 Concordance index comparison



(a) Concordance index comparison of primary biliary cirrhosis dataset (b) Concordance index comparison of German breast cancer dataset

Figure 1

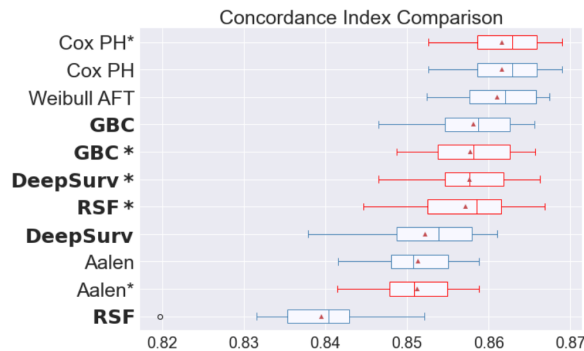


Figure 2: Concordance index comparison of telecom churn dataset

Figure 1a shows the concordance index comparison under the PBC dataset. The methods are shown in decreasing order of their obtained mean score. Note that the mean score value is marked by the red triangle in each box plot. We can observe that

Weibull AFT with the randomized search of the parameters (denoted by Weibull AFT*) is the method that outperforms the others, followed by random survival forest with the randomized search of the parameters (RSF*), random survival forest (RSF) and Cox proportional hazard with the randomized search of the parameters (Cox PH*). We can also see that the randomized search of the parameters works well for all the methods (see Weibull AFT* vs. Weibull AFT, RSF* vs. RSF, Cox PH* vs. Cox PH, GBC* vs. GBC, and Aalen* vs. Aalen, respectively). In particular, Weibull AFT* and Cox PH* obtain an increment of 2.9% and 2.5% against Weibull AFT and Cox PH, respectively.

Figure 1b shows the concordance index comparison result under the GBCSG2 dataset. Here, the method with the best performance is the random survival forest (RSF), followed by gradient boosting Cox (GBC). Unlike the result under the PBC dataset, we cannot observe an increment in the performance when implementing the randomized search of the parameters on RSF, GBC and the other, except for Cox proportional hazard, implementing the randomized search of the parameters (i.e., Cox PH*) has a slight increment of 0.7%. Figure 2 shows the concordance index comparison result under the TLCM dataset. We see that the Cox proportional hazard method (both Cox PH* and Cox PH) outperforms the others, followed by Weibull AFT, whose performance is close to Cox's. In this dataset, we observe that the randomized search does not contribute significantly to improving the performance of the methods, except for the case of random survival forest (RSF) where there is a 2% increment by RSF* when compared with RSF. Weibull AFT* and DeepSurv* were not considered in the graph because they underperformed compared to the other models, and in addition, their performance value was out of the bounds of the figure.

In general, we can observe that for the first two datasets (PBC and GBCSG2), machine learning methods (RSF, RSF*, GBC, GBC* and DeepSurv) perform very well, while parametric methods (Cox PH, Weibull AFT and Aalen additive) are left behind.

This is not the case for the TLCM dataset where Cox HP and Weibull AFT are leading. In addition, we would like to remark the fact that the performance of each method, and its ranking, depends on the dataset. Some methods will perform better

for a certain type of dataset than others, this might be due to the assumptions about the hazard function structure and how these assumptions fit the real distribution of each dataset.

5.2 Integrated Brier score comparison

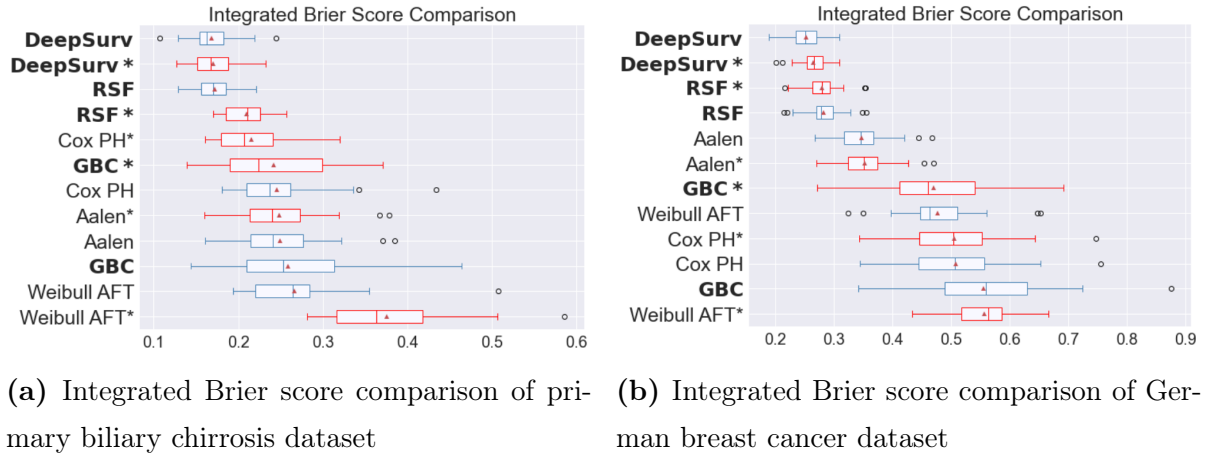


Figure 3

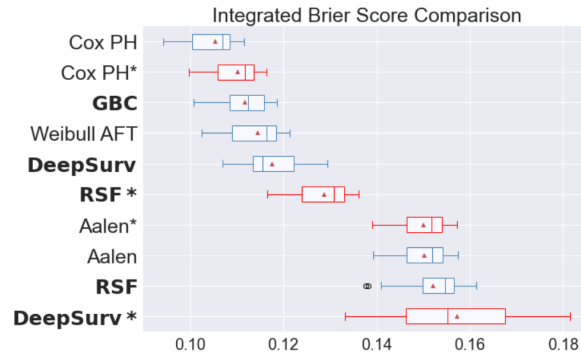


Figure 4: Integrated Brier score comparison of Telecom churn dataset

Figure 3a shows the integrated Brier score comparison under the PBC dataset. In this figure, as in the case of the concordance index, the methods are displayed in the increasing order of performance, which in this case corresponds to decreasing integrated Brier score. Here, we observe that DeepSurv outperforms the other methods (for the IBS score, the lower the better), followed by DeepSurv* and

RSF. We see that there is a clear predominance of machine learning techniques (DeepSurv, DeepSurv*, RSF and RSF*). Similarly, for the GBCSG2 dataset, in Figure 3b, DeepSurv outperforms the other methods, followed by DeepSurv*, RSF*, and RSF. Note that Aalen additive has a performance of 23% worse than that of RSF. In this case, we can also say that machine learning techniques (DeepSurv, DeepSurv*, RSF and RSF*) have better results than the other methods.

Figure 4 shows the integrated Brier score comparison under the TLCM dataset. Contrary to the previous cases, Cox PH method is the lead. In Figure 4, we can appreciate a slight predominance of parametric approaches (Cox PH, Cox PH* and Weibull AFT). We can see that when the amount of censored data is larger, machine learning techniques (DeepSurv and RSF) do not outperform the classical parametric methods.

Finally, we would like to remark that for a given dataset the results for the concordance index and integrated Brier score differ. This is not surprising in this case due to the nature of the two scores, that is very different in between them. Some models can give good ranked results while calibrating very poorly and vice-versa. Discussions about how to choose an appropriate score have taken place in the past and there is no consensus in the community [45].

5.3 Ensemble methods comparison

In the following, we show the result of our deployed ensemble method. Each aggregation is set according to Section 3 for optimizing the parameters of a convex combination of the six methods (described in Section 2.1).

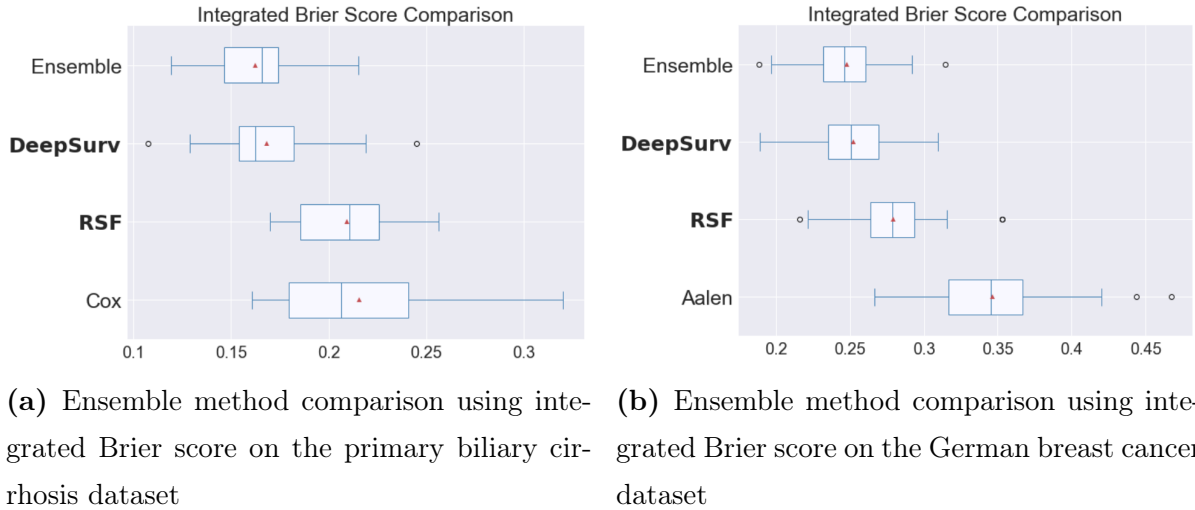


Figure 5

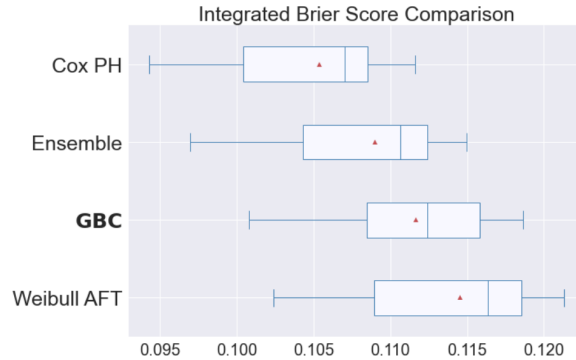


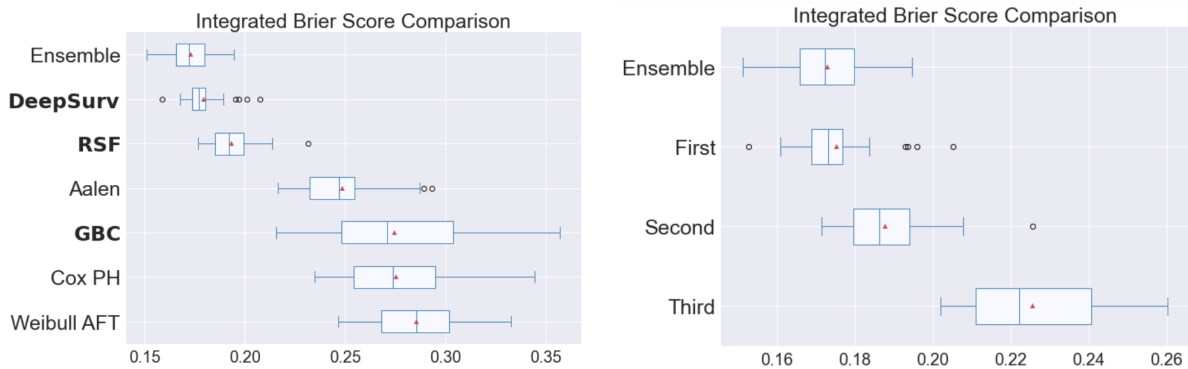
Figure 6: Ensemble method comparison using integrated Brier score on the Telecom churn dataset

Figure 5a shows the integrated Brier score comparison result under the PBC dataset. We observe that the ensemble method through gradient descent outperforms DeepSurv by 3%. Similarly, in Figure 5b for the GBCSG2 dataset, we find that the ensemble method outperforms the best predictor among the six and obtains a performance improvement of 1.6%. Finally, Figure 6 shows the integrated Brier score comparison result under the TLM dataset. The ensemble method does not improve the performance, whereas the best estimator is the Cox PH which has a

performance of 3.8% better than that of the ensemble method.

In addition, we show the overall performance of each method by averaging the scores obtained by each under the three datasets so as to estimate their overall performance.

Figure 7a shows the comparison among all the techniques, including the deployed ensemble method. We see that, in the overall score, the ensemble method outperforms the best predictor by 3.4%. In Figure 7b, we show the result obtained by averaging the best scores obtained by the six methods (described in Section 2.1) in each of the three datasets (they are DeepSurv for PBC and GBCSG2 and Cox PH for TLCM, see Figures 3a, 3b and 4, respectively) to obtain a global score, which corresponds to the average of the best scores among the six algorithms without using the ensemble method. We similarly average the second best scores obtained by the six methods among the three datasets. Finally, this is also applied to the third best scores in the same way. The results are labeled as “First”, “Second” and “Third” in Figure 7b, respectively. We see that the ensemble method improves by 1% the performance of the “First” score and has shown its effectiveness.



(a) Integrated Brier score comparison among the three datasets (b) Integrated Brier score overall comparison among the three datasets

Figure 7

6 SIMULATION EXPERIMENTS

To deepen the insights from Section 5, we conducted experimental simulations with the goal of comparing the ranking of the methods under different dataset

configurations; thus, to understand why some methods perform better than others. We simulate data using three different methods and under three different scenarios. The first two methods were based on R libraries `coxed` [23] and `simsurv` [10]. Both assume a particular shape of the hazard function, Cox proportional hazards and Weibull AFT respectively. The third method was created by us following the logic of the truck dataset from O. Grisel and V. Maladiere [20]. We give a further explanation in Section 6.1. This method was carried out in Python. The presented results are the averages obtained from 100 simulations. We do the comparison using the concordance index, and a similar analysis is presented in Appendix C using the integrated brier score.

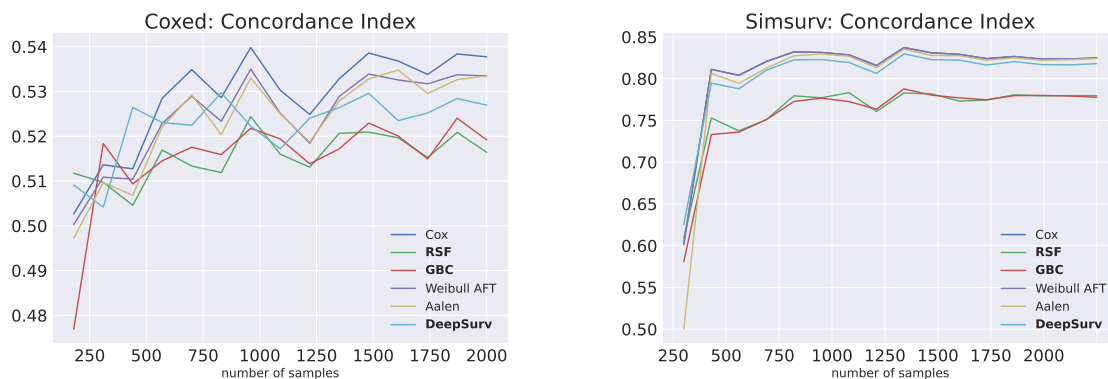
6.1 Python dataset simulation

Following the truck dataset simulation from [20], we first generate a specified number of features d for each of the N individuals. These features include normally distributed $\mathcal{N}(1, 0.3)$ values, uniform $\mathcal{U}(0, 1)$ values and categorical features of 3 categories. Next, we define three types of failure, as mentioned in [20]: initial assembly failure, operation failure and fatigue failure. Although our method aims to be more general than the truck problem, we maintain the distributions specified in the cited reference. Each type of failure is modeled by a different Weibull curve with parameter λ . The first type of failure has a decreasing hazard with $\lambda = 0.003$, while the other two types have hazard rates that increase, with $\lambda = 3$ and $\lambda = 6$, respectively. The influence of the features on each of the failure types will vary in each experiment, depending on the number of features considered. To continue, we sample events of the three types for each individual and we choose the first one that occurs, or none if no event has taken place (censored case). Finally, we incorporate non-informative uniform censoring, where the parameters of the uniform distribution vary for each simulation case. The length of the uniform interval is what provides us with control over the percentage of censorship.

6.2 Number of samples

In this section, we compare the behavior of the methods as the number of samples increases. We consider 12 features and 50% of censorship. We vary the number of samples over a grid in between 50 and 2000 to study the impact of the number of samples in the performance of the different methods. The results are presented in

the following figures.



(a) Concordance index comparison of the increasing sample size simulation with Coxed library

(b) Concordance index comparison of the increasing sample size simulation with Simsurv library

Figure 8

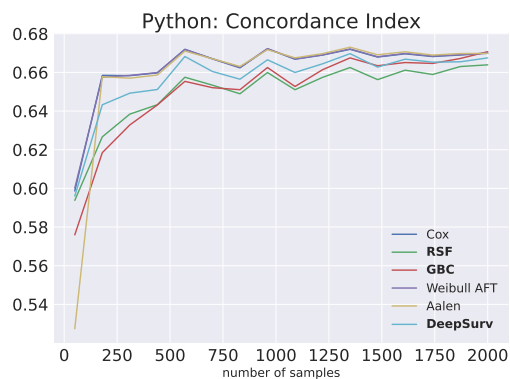


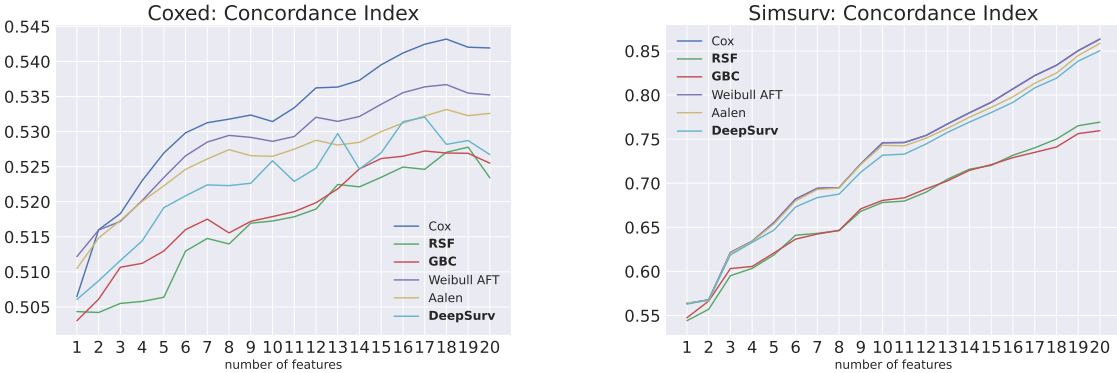
Figure 9: Concordance index comparison of the increasing sample size simulation with Python

We observe in Figures 8a, 8b, and 9 that the concordance index improves as the number of samples increases. Additionally, we observe in Figure 8a that Cox proportional hazard consistently outperforms the other methods, regardless of the

number of samples. Subsequently, the order is not very clear, but random survival forest and gradient boosting consistently show lower performance. In Figures 8b and 9, we observe a consistent outperformance of Cox proportional hazards and Weibull AFT, closely followed by Aalen additive hazards. Random survival forest and gradient boosting underperform compared to the other methods in both figures. In conclusion, the ranking of the models performance appears to depend on the shape of the underlying distribution used to sample the event times and not on the number of samples of the dataset.

6.3 Number of features

In this section, we compare the behavior of the methods as the number of features decreases. With a fixed 50% of censorship and 1000 samples, we start the analysis with 20 features and progressively remove one feature at each step. The results are presented in the following figures.



(a) Concordance index comparison of the decreasing number of features simulation with Coxed library (b) Concordance index comparison of the decreasing number of features simulation with Simsurv library

Figure 10

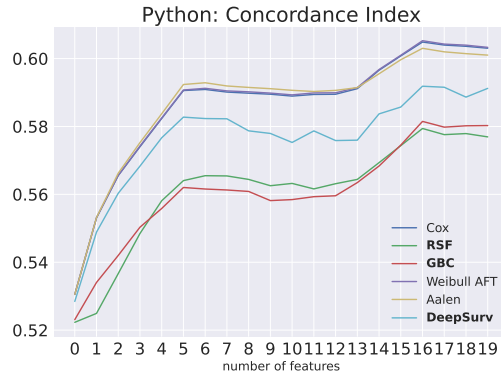


Figure 11: Concordance index comparison of the decreasing number of features simulation with Python

We note in Figures 10a, 10b, and 11 that the concordance index improves as the number of features increases. This behavior aligns with our expectations since the initial model is constructed with 20 features, and the subsequent removal of features results in a reduction of information. Moreover, we observe in Figure 10a that, as in Figure 8a, Cox proportional hazard consistently outperforms the other methods, regardless of the number of features. The same holds for Figure 10b and 11, where the best performance is shared by Cox proportional hazard, Weibull AFT, and Aalen additive hazards. Following the conclusion of Section 6.2, the ranking of the models depends mainly on the shape of the distribution used to generate the data, rather than on the number of features.

6.4 Percentage of censorship

In this section, we compare the behavior of the methods as the percentage of censorship increases. We fix the number of samples at 1000 and the number of features at 12. The results are presented in the following figures.



(a) Concordance index comparison of the increasing percentage of censorship simulation with Coxed library (b) Concordance index comparison of the increasing percentage of censorship simulation with Simsurv library

Figure 12

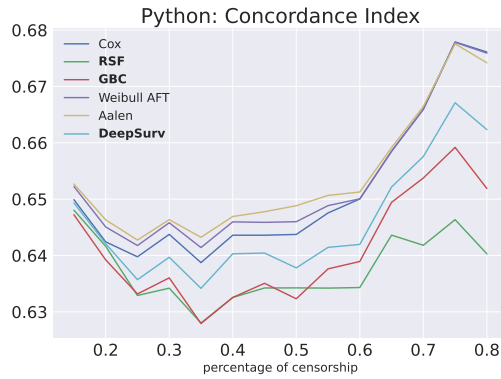


Figure 13: Concordance index comparison of the increasing percentage of censorship simulation with Python

In Figure 12a, we observe a decline in performance as the percentage of censorship increases. This is in line with the notion that higher levels of censorship result in reduced available information, consequently leading to diminished performance. However, a contrasting pattern emerges in Figure 12b and 13, where we actually observe an improvement in performance towards the end of the curves. We believe

that this phenomenon is attributed to a bias in the concordance index when the percentage of censorship is high. One solution to address this issue is presented by Uno et al. [44], where they introduced a weighted version of the score. In addition, we observe in Figure 12a that Cox proportional hazards outperforms the other methods, followed by Weibull AFT and Aalen additive. This same pattern is evident in Figures 12b and 13, where these three models lead in terms of performance. Notably, the ranking of the methods remains consistent even as the percentage of censorship increases, reinforcing the conclusion from the previous sections. The primary factor influencing the performance change of the methods is the congruence between the model assumptions and the actual distribution of event times, with improved fit leading to better performance.

7 CONCLUSIONS

This paper presents an extensive analysis of different survival methods applied to three datasets and compared by two scoring rules. The study shows how diverse a single method’s performance is when changing the measure of comparison and when it is applied to datasets of different distributions, sizes and percentages of censorship. We propose a straightforward aggregation of methods of different natures, parametric, semi-parametric and machine learning, that assume diverse shapes of the hazard function allowing the ensemble model to gain in robustness with respect to each single predictor. This can be observed in Figure 7a by the outperformance of the assemblage measured by an overall score that is independent of the dataset. Finally, we present simulation experiments with the objective of studying which dataset characteristics have the most significant influence on the performance of the models. This analysis leads us to the conclusion that the proximity of the model assumptions to the real event distribution is a determining factor in performance. Further research could go in the direction of complexifying the combination algorithm by considering time-varying weightings and more sophisticated optimization procedures. Another direction could be to find theoretical guarantees for the integrated Brier score of the ensemble method and possibly in a stochastic setting.

REFERENCES

- [1] Odd O Aalen. A linear regression model for the analysis of life times. *Statistics in Medicine*, 8(8):907–925, 1989.
- [2] Odd O Aalen and Thomas H Scheike. Aalen’s additive regression model. *Encyclopedia of Biostatistics*, 1, 2005.
- [3] Sercan O Arık and Tomas Pfister. Tabnet: Attentive interpretable tabular learning. *arXiv*, 2020.
- [4] Jonathan Berrisch and Florian Ziel. Crps learning. *Journal of Econometrics*, 2021.
- [5] Harald Binder and Martin Schumacher. Allowing for mandatory covariates in boosting estimation of sparse high-dimensional survival models. *BMC Bioinformatics*, 9(1):1–10, 2008.
- [6] J Martin Bland and Douglas G Altman. The logrank test. *British Medical Journal (BMJ)*, 328(7447):1073, 2004.
- [7] Leo Breiman. Pasting bites together for prediction in large data sets and on-line. *Univ. of Calif., Berkeley, Dept. of Statistics Technical Report*, 1997.
- [8] Leo Breiman. Random forests. *Machine Learning*, 45(1):5–32, 2001.
- [9] Glenn W Brier et al. Verification of forecasts expressed in terms of probability. *Monthly Weather Review*, 78(1):1–3, 1950.
- [10] Samuel L Brilleman, Rory Wolfe, Margarita Moreno-Betancur, and Michael J Crowther. Simulating survival data using the `simsurv` r package. *Journal of Statistical Software*, 97:1–27, 2021.
- [11] David R Cox. Regression models and life-tables. *Journal of the Royal Statistical Society: Series B (Methodological)*, 34(2):187–202, 1972.
- [12] Sidney J Cutler and Fred Ederer. Maximum utilization of the life table method in analyzing survival. *Journal of Chronic Diseases*, 8(6):699–712, 1958.

- [13] Cameron Davidson-Pilon. Lifelines: survival analysis in Python. *Journal of Open Source Software*, 4(40):1317, 2019.
- [14] Thomas G Dietterich. Ensemble methods in machine learning. In *International Workshop on Multiple Classifier Systems*, pages 1–15. Springer, 2000.
- [15] Stephane Fotso et al. PySurvival: Open source package for survival analysis modeling, 2019–. URL <https://www.pysurvival.io/>.
- [16] Jerome H Friedman. Greedy function approximation: a gradient boosting machine. *Annals of Statistics*, pages 1189–1232, 2001.
- [17] Nir Friedman, Dan Geiger, and Moises Goldszmidt. Bayesian network classifiers. *Machine Learning*, 29(2):131–163, 1997.
- [18] Thomas A Gerds and Martin Schumacher. Consistent estimation of the expected brier score in general survival models with right-censored event times. *Biometrical Journal*, 48(6):1029–1040, 2006.
- [19] Erika Graf, Claudia Schmoor, Willi Sauerbrei, and Martin Schumacher. Assessment and comparison of prognostic classification schemes for survival data. *Statistics in Medicine*, 18(17-18):2529–2545, 1999.
- [20] O. Grisel and V. Maladiere. Survival analysis benchmark. https://github.com/soda-inria/survival-analysis-benchmark/blob/main/notebooks/truck_dataset.ipynb, 2023.
- [21] Humza Haider, Bret Hoehn, Sarah Davis, and Russell Greiner. Effective ways to build and evaluate individual survival distributions. *J. Mach. Learn. Res.*, 21(85):1–63, 2020.
- [22] Lars Kai Hansen and Peter Salamon. Neural network ensembles. *IEEE Transactions on Pattern Analysis and Machine Intelligence*, 12(10):993–1001, 1990.
- [23] Jeffrey J Harden and Jonathan Kropko. Simulating duration data for the cox model. *Political Science Research and Methods*, 7(4):921–928, 2019.

- [24] Frank E Harrell Jr, Kerry L Lee, and Daniel B Mark. Multivariable prognostic models: issues in developing models, evaluating assumptions and adequacy, and measuring and reducing errors. *Statistics in Medicine*, 15(4):361–387, 1996.
- [25] Trevor Hastie, Robert Tibshirani, Jerome H Friedman, and Jerome H Friedman. *The Elements of Statistical Learning: Data Mining, Inference, and Prediction*, volume 2. Springer, Berlin, Germany, 2009.
- [26] IBM. Kaggle telco customer churn:. *IBM Cognos Analytics*, 2008.
- [27] Hemant Ishwaran, Udaya B Kogalur, Eugene H Blackstone, and Michael S Lauer. Random survival forests. *The Annals of Applied Statistics*, 2(3):841–860, 2008.
- [28] Edward L Kaplan and Paul Meier. Nonparametric estimation from incomplete observations. *Journal of the American Statistical Association*, 53(282):457–481, 1958.
- [29] Jared L Katzman, Uri Shaham, Alexander Cloninger, Jonathan Bates, Tingting Jiang, and Yuval Kluger. DeepSurv: personalized treatment recommender system using a Cox proportional hazards deep neural network. *BMC Medical Research Methodology*, 18(1):1–12, 2018.
- [30] Faisal M Khan and Valentina Bayer Zubek. Support vector regression for censored data (SVRc): a novel tool for survival analysis. In *IEEE International Conference on Data Mining*, pages 863–868. IEEE, 2008.
- [31] William R Lane, Stephen W Looney, and James W Wansley. An application of the Cox proportional hazards model to bank failure. *Journal of Banking & Finance*, 10(4):511–531, 1986.
- [32] Jerald F Lawless and Yan Yuan. Estimation of prediction error for survival models. *Statistics in Medicine*, 29(2):262–274, 2010.

- [33] Kung-Yee Liang, Steven G Self, and Xinhua Liu. The Cox proportional hazards model with change point: An epidemiologic application. *Biometrics*, pages 783–793, 1990.
- [34] Sebastian Pölsterl. scikit-survival: A library for time-to-event analysis built on top of scikit-learn. *J. Mach. Learn. Res.*, 21(212):1–6, 2020.
- [35] Sebastian Pölsterl, Nassir Navab, and Amin Katouzian. Fast training of support vector machines for survival analysis. In *Joint European Conference on Machine Learning and Knowledge Discovery in Databases*, pages 243–259. Springer, 2015.
- [36] Greg Ridgeway. The state of boosting. *Computing science and statistics*, pages 172–181, 1999.
- [37] Sebastian Ruder. An overview of gradient descent optimization algorithms. *arXiv preprint arXiv:1609.04747*, 2016.
- [38] Willi Sauerbrei and Martin Schumacher. A bootstrap resampling procedure for model building: application to the Cox regression model. *Statistics in Medicine*, 11(16):2093–2109, 1992.
- [39] Robert E Schapire. The strength of weak learnability. *Machine Learning*, 5(2):197–227, 1990.
- [40] M. Schumacher, G. Bastert, H. Bojar, K. Hübner, M. Olschewski, W. Sauerbrei, C. Schmoor, C. Beyerle, R. L Neumann, and H. F Rauschecker. Randomized 2 × 2 trial evaluating hormonal treatment and the duration of chemotherapy in node-positive breast cancer patients. *German Breast Cancer Study Group. Journal of Clinical Oncology*, 12(10):2086–2093, 1994. URL <https://www.pysurvival.io/>.
- [41] Harald Steck, Balaji Krishnapuram, Cary Dehing-Oberije, Philippe Lambin, and Vikas C Raykar. On ranking in survival analysis: Bounds on the concordance index. In *Advances in Neural Information Processing Systems*, pages 1209–1216, 2008.

- [42] G. Therneau, T. Grambsch. Modeling Survival Data: Extending the Cox Model. *Springer-Verlag*, 2000.
- [43] James Tobin. Estimation of relationships for limited dependent variables. *Econometrica: Journal of the Econometric Society*, pages 24–36, 1958.
- [44] Hajime Uno, Tianxi Cai, Michael J Pencina, Ralph B D’Agostino, and Lee-Jen Wei. On the c-statistics for evaluating overall adequacy of risk prediction procedures with censored survival data. *Statistics in medicine*, 30(10):1105–1117, 2011.
- [45] Wessel N Van Wieringen, David Kun, Regina Hampel, and Anne-Laure Boulesteix. Survival prediction using gene expression data: a review and comparison. *Computational Statistics & Data Analysis*, 53(5):1590–1603, 2009.
- [46] Ping Wang, Yan Li, and Chandan K Reddy. Machine learning for survival analysis: A survey. *ACM Computing Surveys (CSUR)*, 51(6):1–36, 2019.
- [47] Waloddi Weibull. A statistical theory of strength of materials. *IVB-Handl.*, 1939.
- [48] Cha Zhang and Yunqian Ma. *Ensemble Machine Learning: Methods and Applications*. Springer, Berlin, Germany, 2012.
- [49] Zhi-Hua Zhou. *Ensemble Methods: Foundations and Algorithms*. Chapman and Hall/CRC, London, UK, 2019.
- [50] Joachim Zietz, Emily Norman Zietz, and G Stacy Sirmans. Determinants of house prices: a quantile regression approach. *The Journal of Real Estate Finance and Economics*, 37(4):317–333, 2008.

A SCORING RULES

A.1 Concordance Index

The concordance index was introduced by Harrell et al. [24] and it is the most widely used performance metric for time-to-event analysis [41]. It measures the fraction of pairs of subjects that are correctly ordered within all the possible pairs that can be ordered. The highest (and best) value that can be obtained is 1, which

means that there is a complete agreement between the order of the observed and predicted times. The lowest value that can be obtained is 0, which means that all the prediction pairs are ordered backward with respect to the observed times, while a value of 0.5 denotes a random model.

First, we take every pair in the test set such that the earlier observed time is not censored. Then, we consider only pairs (i, j) such that $i < j$ and we also eliminate the pairs for which the times are tied. Next, we define a score $C_{i,j}$ for each pair (i, j) such as $y_i \neq y_j$, equal to 1 if the subject with earlier time (between i and j) has higher predicted risk, equal to 0.5 if the risks are tied, or equal to 0 otherwise.

Finally, given a subset of the data \mathcal{D} of size n , we compute the concordance index as follows:

$$CI(\hat{S}, \mathcal{D}) = \frac{1}{|\mathcal{P}|} \sum_{(i,j) \in \mathcal{P}} C_{i,j},$$

where,

$$C_{i,j} = \begin{cases} 1 & \text{if } y_i < y_j \text{ and } \hat{R}(x_i) > \hat{R}(x_j) \\ 0.5 & \text{if } \hat{R}(x_i) = \hat{R}(x_j) \\ 0 & \text{otherwise,} \end{cases}$$

and $\mathcal{P} = \{(i, j) \in \mathcal{D} \times \mathcal{D} : i < j, y_i \neq y_j, \text{ if } y_i < y_j, \text{ then } \delta_i = 1\}$ is the set of all eligible pairs. To calculate the concordance index, we use the version of scikit-survival library [34] in Python.

A.2 Integrated Brier score

We consider an approach based on the estimates of the probability functions that will be used as predictions of the event status $\mathbb{1}\{T_i > t\}$. In this case, $\mathbb{1}\{T_i > t\}$ has to be compared with $\hat{S}(t|X_i)$, leading to the mean squared error (*MSE*) at time t :

$$MSE(\hat{S}, t) = \mathbb{E}[(\mathbb{1}\{T_i > t\} - \hat{S}(t|X_i))^2].$$

The Brier score, introduced initially to measure the inaccuracy of probabilistic weather forecast by Brier [9], is an estimator of the *MSE*. It is important to remark that the *MSE* cannot be directly computed from the dataset since we do not know the underlying distribution of T_i but only the realizations of Y_i . Let us

define $S_C(t|X_i) = \mathbb{P}(C_i > t|X_i)$ the survival censoring distribution and the Brier score:

$$BS(\hat{S}, t, \mathcal{D}) = \frac{1}{n} \sum_{i=1}^n W_i(t) (\mathbb{1}\{y_i > t\} - \hat{S}(t|x_i))^2,$$

where (x_i, y_i, δ_i) for $0 < i \leq n$ are points from \mathcal{D} , and W_i is defined for all t as:

$$W_i(t) = \frac{\delta_i \mathbb{1}\{y_i \leq t\}}{\hat{S}_C(y_i|x_i)} + \frac{\mathbb{1}\{y_i > t\}}{\hat{S}_C(t|x_i)}.$$

Gerds and Schumacher [18] proved that the Brier score is a consistent estimator for the mean square error when the estimation \hat{S}_C of S_C is well specified. Let us notice that in our implementation of the score, we use a Kaplan-Meier [28] estimator for the survival censoring function \hat{S}_C , which does not depend on the covariates. This assumption is not always the case for the real censoring function S_C , and it can lead to misspecifications of the model (wrong hypothesis on the probability shape) and, thus, to an estimation bias [19].

Finally, we consider over this paper the integrated Brier score:

$$IBS(\hat{S}, \mathcal{D}) = \frac{1}{\tau} \int_0^{\tau} BS(\hat{S}, t, \mathcal{D}) dt,$$

where τ is a user-specified time horizon. There exist diverse scoring rules for survival models based on $L1$ -loss, logarithmic loss and 1-calibration in between others (see [21] and [19] for more details). Other approaches of the estimation of prediction errors and model misspecification can be found in [32]. We chose the concordance index and integrated Brier score because they measure different aspects of the models, ranking and calibration, allowing us to have a good understanding of the performance of the methods.

B IMPLEMENTED METHODS

B.1 Cox Proportional Hazard (Cox PH)

Cox proportional hazard is a semi-parametric method proposed by Cox [11] with the objective of measuring the impact of each covariate/feature in the estimation of the survival probability function. It models the hazard function as a general linear regression of the covariates and a non-parametric baseline function $\lambda_0(t)$ that

depends only on time. Given a subject with a covariate vector $x = \{x^1, \dots, x^d\}$, the hazard function is as follows:

$$h(t|x) = \lambda_0(t) \exp(\beta^T x),$$

where the parameter $\beta = (\beta_1, \dots, \beta_d)$ is estimated by maximizing the likelihood. This model is semi-parametric in the sense that the baseline function $\lambda_0(t)$ does not need to be specified and it can be chosen differently for each unique time. Cox proportional hazard is one of the most often used methods in time-to-event analysis and has a wide range of applications [31], [33], [38]. We use the implementation from scikit-survival library [34], where a regularization parameter α for ridge regression penalty is used and whose default is equal to 0. The mortality risk prediction will be determined by the log hazard ratio $R(x) = \beta^T x$.

B.2 Gradient Boosting Cox (GBC)

Gradient boosting Cox is a machine learning method that was first proposed by Breiman [7], developed by Friedman [16] and adapted to survival models by Ridgeway [36]. The main idea is to combine a series of base learners in an additive manner to obtain a strong overall model. The base learners implemented in this case will be regression trees fitted at each stage on the negative gradient of the loss function. This is an additive method in the sense that it is constructed sequentially in a step-by-step greedy way. We can define the overall function f as follows:

$$f(x) = \sum_{k=1}^K \rho_k \cdot g_k(x, \theta),$$

where g_k is used to denote the base learners and K is the number of learners. Therefore, the objective is to maximize the log-likelihood function of Cox's proportional hazard model by replacing the linear regression $\beta^T x$ with the additive function $f(x)$ such that we have the following expression for the hazard function:

$$h(t|x) = \lambda_0(t) \exp(f(x)).$$

We use the implementation of scikit-survival [34] where we find three parameters of our interest, the learning rate that shrinks the contribution of each tree and it is set as default by 0.1, the maximum depth that specifies the depth to which each tree

will be built and that is set equal to 3 by default, and the minimum samples leaf that determines the number of samples required to be at a leaf node and its default is equal to 1. Similar as Cox proportional hazard the mortality risk prediction can be interpreted as the log hazard ratio $f(x)$.

B.3 Random Survival Forest (RSF)

Random survival forest was proposed by Ishwaran et al. [27] as an adaptation for censored data of the random forest method introduced by Breiman et al. [8]. It is an ensemble of tree-based learners where each tree is built from a bootstrap sampling of the training set in order to reduce the correlation between the trees. Also, for each node, it only evaluates the split criterion for a random subset of features and thresholds. The quality of a split is measured by the log-rank splitting rule [6] and then predictions are formed by aggregating predictions of the individual trees.

We implemented random survival forest from scikit-survival [34] and we will consider three of its parameters. The first one is the maximum depth which is set as infinity, which means that the nodes are expanded until no further partitioning is possible. The second one is the maximum features number which indicates the maximum number of features to consider when looking for the best split; this parameter is set as the number of data features. The last parameter is the minimum samples leaf which in this case the default value is 3. Here, the mortality risk is defined by the ensemble mortality (see [27] for details) which corresponds to the sum of the cumulative Hazard functions estimated by the forest.

B.4 Weibull Accelerated Failure Time (Weibull AFT)

Weibull AFT is a parametric model that was named after Waloddi Weibull, who was the first to promote its usefulness, particularly in the domain of strength of materials [47]. Accelerated failure time models also assume that the effect of a covariate is to accelerate or decelerate the life course. Given the parameters ρ and λ , the survival function of the Weibull distribution can be given as:

$$S(t|x) = \exp\left(-\left(\frac{t}{\rho(x)}\right)^\lambda\right),$$

where we consider the scale parameter $\rho(x) = \exp(\beta_0 \cdot (\beta^T x))$ and λ is the parameter that controls the concavity of the cumulative hazard, indicating acceleration or

deceleration hazards. In this case, we implement Weibull AFT from lifelines library [13] and we consider a penalizer parameter and a ℓ_1 -ratio to adjust how much of the penalizer should be attributed to an ℓ_1 penalty. Both of them are initially set as zero by default. Here, the mortality risk is defined by $\mathbb{E}[T_i|x_i]$.

B.5 Aalen’s Additive Fitter (Aalen)

Aalen’s additive is a parametric method proposed by Aalen [1]. This model responds to the fact that not all the covariates effects must be proportional, which is different from the assumption of Cox proportional hazard, but some of them can have additive effects. Besides, Aalen’s additive model allows the effects of the covariates to vary over time which is not always the case with the other methods. The hazard function, in this case, is given as follows:

$$h(t|x) = \beta_0(t) + \beta^T(t)x,$$

where $\beta(t)$ is an unknown parameter of dimension d that are estimated by a linear regression (see [2]). We consider only the penalizer coefficient, which attaches an ℓ_2 penalizer to the size of the parameters during regression that improves the stability of the estimations and controls the high correlation of the features. This penalizer is set to zero by default. Similarly as Weibull AFT, the mortality risk is defined by $\mathbb{E}[T_i|x_i]$.

B.6 DeepSurv

DeepSurv is a nonlinear version of the Cox proportional hazard method proposed by Katzman et al. [29]. DeepSurv allows the use of neural networks within the original design of Cox’s and aims to offer more flexibility in terms of the structure of the model than Cox proportional hazard. DeepSurv is a multi-layer perceptron that predicts the risk of failure. The output of the network $\hat{r}_\theta(x)$ is a single node that estimates the risk function. The loss function to minimize is the negative log-partial likelihood of the Cox proportional hazard method:

$$\begin{aligned} \ell(\theta) = & -\frac{1}{L} \sum_{i:\delta_i=1} \left(\hat{r}_\theta(x_i) - \log \left(\sum_{j \in \mathcal{R}(T_i)} e^{\hat{r}_\theta(x_j)} \right) \right) \\ & + \lambda \|\theta\|^2, \end{aligned}$$

where λ is a ℓ_2 regularization parameter and L is the number of uncensored subjects. The network weights that minimize the loss function can be estimated by a

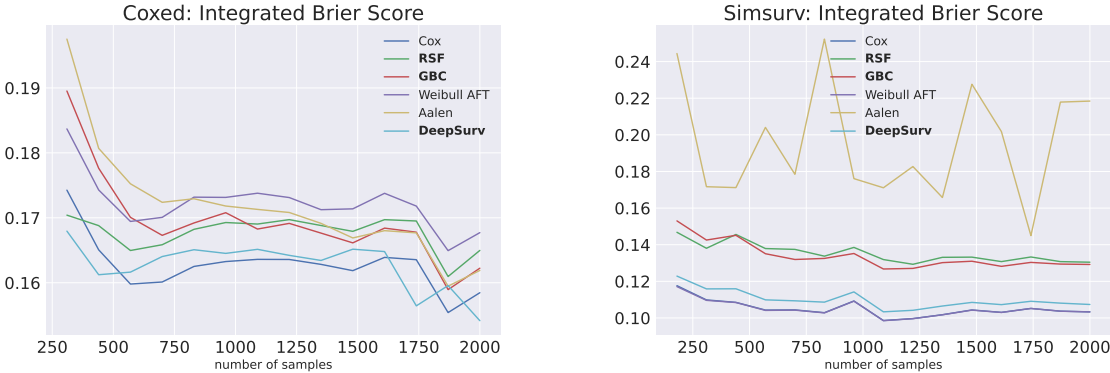
gradient descent algorithm [37]. We use the implementation from Pysurvival [15], where we can choose the structure of the multilayer perceptron by choosing the number of hidden units per layer. We will consider two fully connected hidden layers and, consequently, two parameters to be set. The default number of units is 60 for the first layer and 10 for the second.

C SIMULATION EXPERIMENTS

In this section, we simulate data using three different techniques. The first set of events is generated by sampling a Cox proportional hazards model, the second by following a Weibull distribution, and the third involves a combination of Weibull distributions. The objective is to compare how the ranking of the methods varies across three experiments. Thus, to understand how different data characteristics can impact the performance of the methods. These findings align with those presented in Section 6, with the distinction that we assess performance using the integrated Brier score.

C.1 Number of samples

The first experiment consists on evaluating the performance of the models as the number of samples increases from 50 to 2000.



(a) Integrated Brier score comparison of the increasing sample size simulation with Coxed library (b) Integrated Brier score comparison of the increasing sample size simulation with Simsurv library

Figure 14

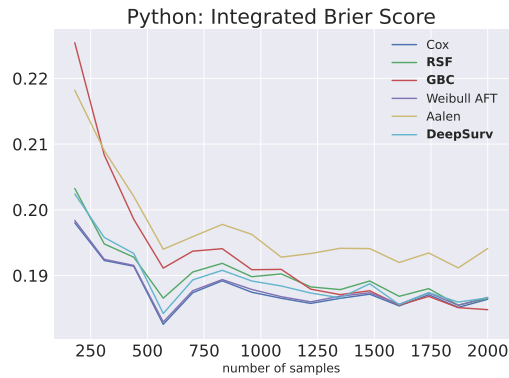
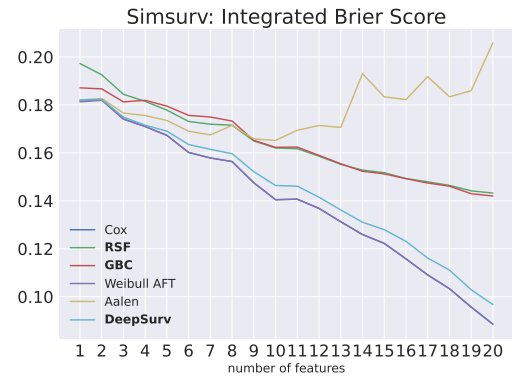
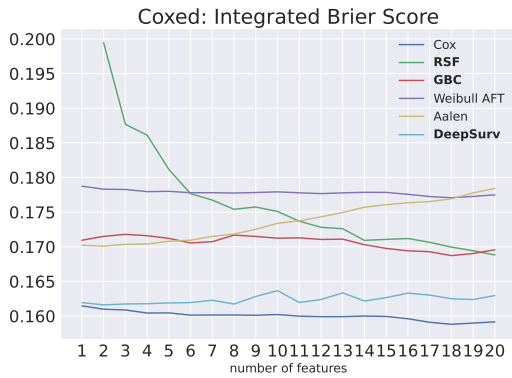


Figure 15: Integrated Brier score comparison of the increasing sample size simulation with Python

We observe in Figures 14a, 14b, and 15, as discussed in Section 6.2, that the performance improves as the number of samples increases. Furthermore, it is noteworthy that the hierarchy of the models remains relatively stable as the number of samples increases. Specifically, Cox proportional hazards outperforms the other methods, with DeepSurv as the second-best performer. This reaffirms the conclusion made in Section 6.2 that the models’ performance order is independent of the sample size but instead depends on the matching between the underlying assumptions and the dataset real distribution shape.

C.2 Number of features

The second experiment consists on evaluating the performance of the models as the number of features decreases from 20 to 1.



(a) Integrated Brier score comparison of the decreasing number of features simulation with Coxed library

(b) Integrated Brier score comparison of the decreasing number of features simulation with Simsurv library

Figure 16

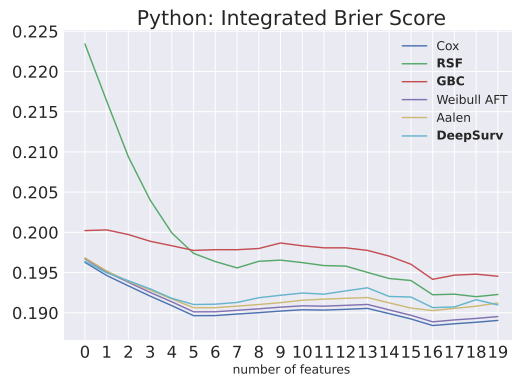


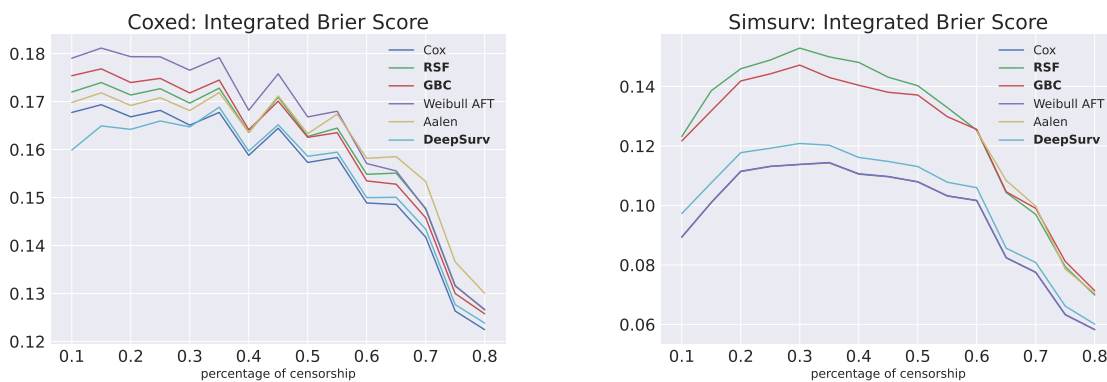
Figure 17: Integrated Brier score comparison of the decreasing number of features simulation with Python

We observe in Figures 16a, 16b and 17 a slight increase in performance as the number of features increases. The most significant change occurs with Random Survival Forest, which exhibits very poor performance compared to the other methods when the number of features is small but becomes competitive as the number of features increases. This is because random survival forest relies on the diversity and richness

of features to make accurate predictions. In addition, in Figures 16a and 16b, we observe that the Aalen additive model does not align with the trend of the other methods, as its performance worsens with an increasing number of features. This could be due to the challenge posed by the additive linearity assumption in capturing the true underlying relationship between covariates and survival outcomes. Finally, we observe that the performance of the Weibull AFT model improves relatively in Figures 16b and 17 compared to Figure 16a. This phenomenon occurs because both simulations, those conducted by the `simsurv` library and our method implemented in python, are based on the Weibull distribution.

C.3 Percentage of censorship

The third experiment, as presented in Section 6, consists on increasing the percentage of censorship from 10% to 80%.



(a) Integrated Brier score comparison of the increasing percentage of censorship simulation with Coxed library (b) Integrated Brier score comparison of the increasing percentage of censorship simulation with Simsurv library

Figure 18

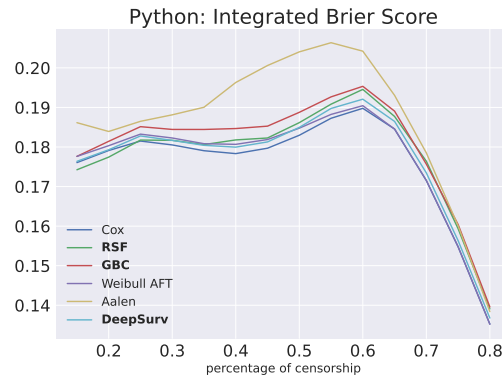


Figure 19: Integrated Brier score comparison of the increasing percentage of censorship simulation with Python

We observe in Figures 18b and 19 an irregular increase in IBS up to 30% and 60% of censorship, respectively. This corresponds to the intuitive expectation that higher censorship rates should lead to poorer performance. However, when the percentage of censorship is high, as can also be seen across the entire curve in Figure 18a, we observe an improvement in performance. This phenomenon might be attributed to the distribution of censorship. Censored individuals contribute to the Brier score only until their observed time. Therefore, if their observed time occurs at the beginning of the observation period, their contribution to the score is minimal. Consequently, if there is a significant percentage of censorship, the Brier score risks being small. Additionally, we observe that there is no significant variation in the hierarchy of the models. In Figure 18a, Cox PH and DeepSurv consistently maintain the lead throughout the entire experiment, while in Figures 18b and 19, Cox PH and Weibull AFT remain at the forefront. This corroborates the conclusion drawn in the previous sections, where we found that the main factor determining the ranking of performance is the underlying distribution of the data, rather than the size of the dataset or the percentage of censorship.