

Infering meteorological information at different scales from several sources of data

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Infering meteorological information at different scales from several sources of data

Didier Josselin*, Matthieu Vignal*, Nicolas Viaux*, Delphine Blanke**, Céline Lacaux**



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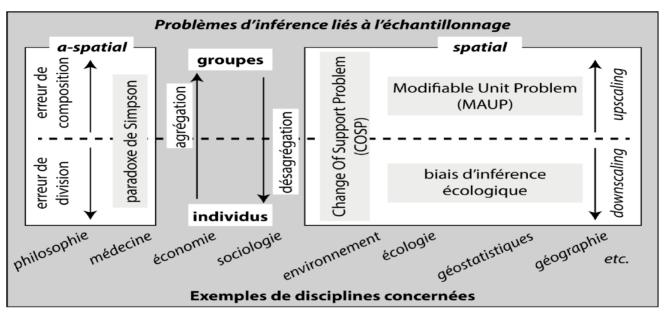


Outline

- Effect of the support on spatial statistics
- Data aggregation through scales
- Eliminating the spatial support effect by resampling and Relative Scalar Deviation calculation
- Conclusion

Effect of the spatial support

Spatial (dis)aggregation



[Openshaw, 1974]

[Yule, 1911 Theil, 1972] [King et al., 2004, Josselin et al., 2004] [Robinson, 1950, Goodman, 1953, King, 1997]

Josselin & Louvet, 2016









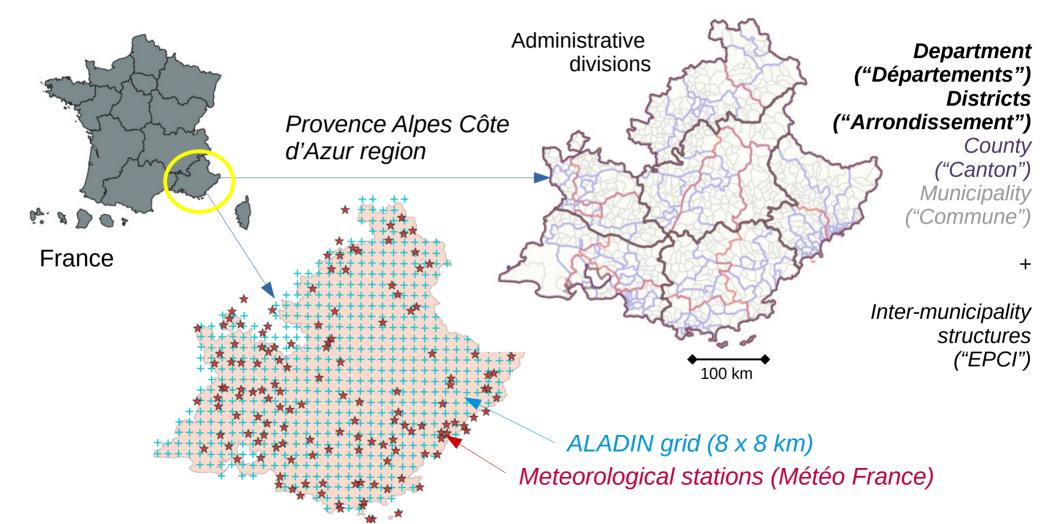


Objectives of the research

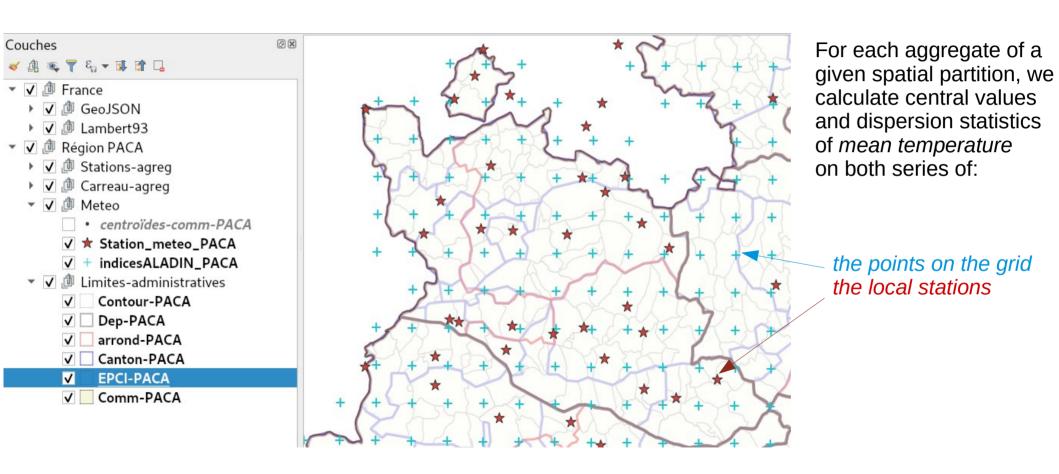
- Studying the relationship between climatic statistical data and aggregation scales
- Finding an accurate and relevant scale for climatic data, at a given level (*i.e.* administrative division)
- Generalizing a method of resampling to eliminate the spatial support effect in rescaling procedures

Data aggregation through scales

Data used (over 30 years)

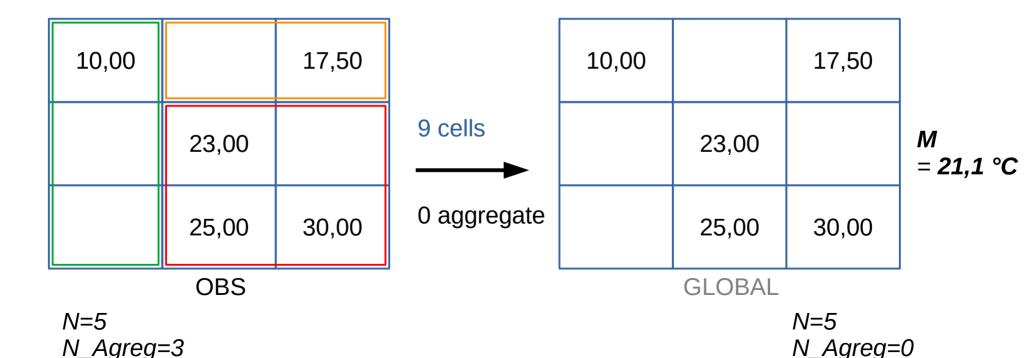


Aggregation process



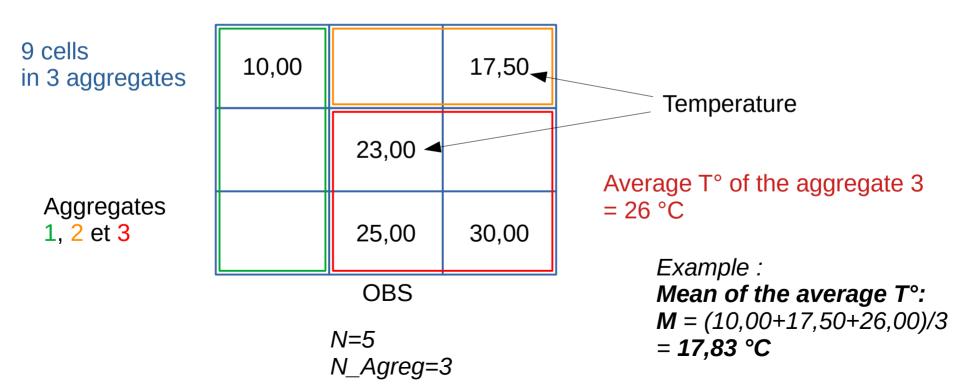
Calculation without partition (e.g. global average M)

No resampling, we compute the clue on observed data without spatial partition

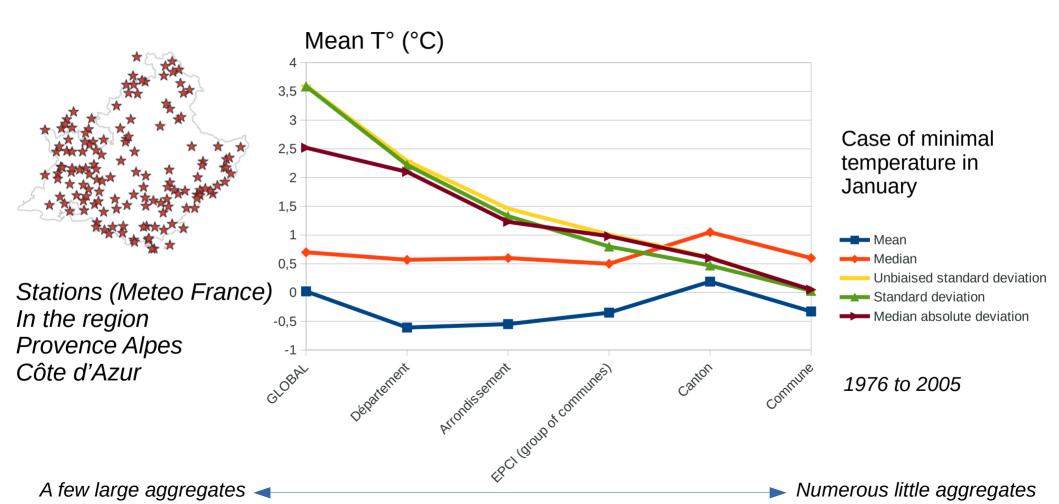


Calculation considering the partition (e.g. aggregated average M)

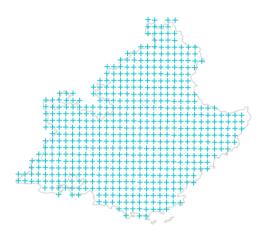
We compute the clue on the observed data for each aggregate, and then aggregate them



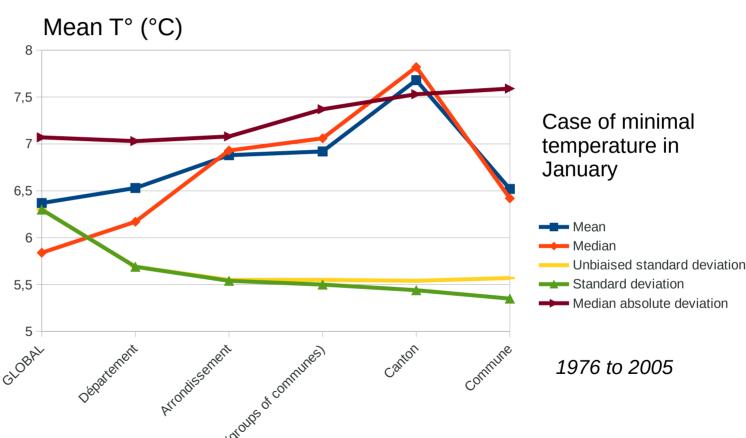
Scalogram of temperatures



Scalogram of temperatures



ALADIN climatic model from CRNM in region Provence Alpes Côte d'Azur



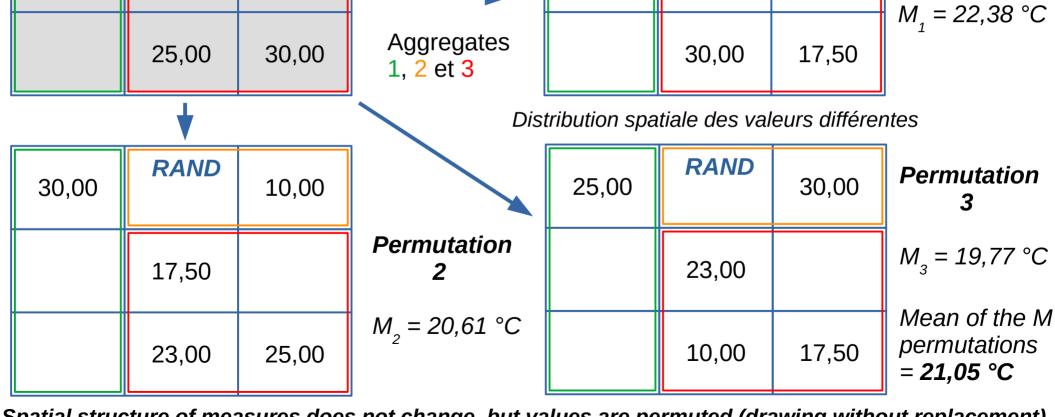
Observations

- There are significant gaps of mean temperatures according to partitions (up to 3.5 or 8 °C)
- Statistical dispersion decreases when number of aggregates increase
- We observe a peak where mean T° is maximal
- But: we need a reference to allow comparison between partitions

Eliminating the spatial support effect by resampling and Relative Deviation calculation (Josselin et al., 2012, 2023)

Resampling procedure

- We randomly permute N times the observed temperatures without changing the spatial partition
- We re-compute statistical clues for each partition
- This is our "control tube" that draws a random spatial distribution of the temperature, without spatial autocorrelation any more



N=5

9 cells

N_Agreg=3

RAND

10,00

25,00

Permutation

23,00

OBS

23,00

17,50

10,00

Spatial structure of measures does not change, but values are permuted (drawing without replacement)

Relative Scalar Deviation

$$RSD(\%) = 100 * \frac{T_{obs} - T_{rand}}{T_{rand}}$$

- Due to permutations, the normalized difference between the observed and the randomized clue allows to eliminate the change of support problem because its effect is similar in both cases (for a given scale/partition)
- The Relative Scalar Deviation reflects the effective part of the geography in the measured values (e.g. Temperature), because the random process deleted all the spatial autocorrelation

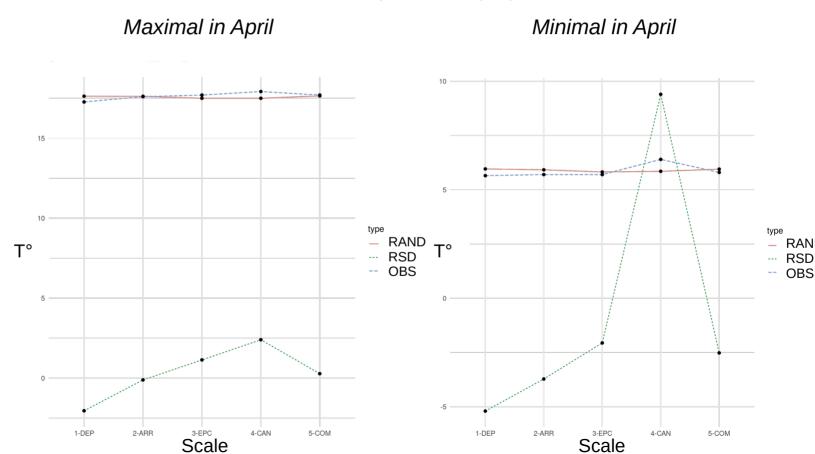
Scalogram with Relative Scalar Deviation





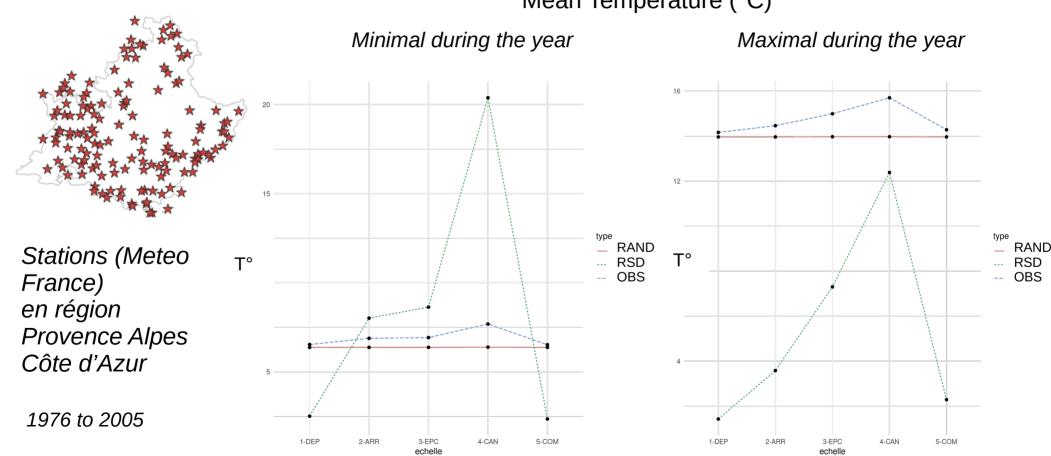
ALADIN climatic model from CRNM in region Provence Alpes Côte d'Azur

1976 to 2005



Scalogram with Relative Scalar Deviation





Conclusion

- A proposition to eliminate the Change of Spatial Support Problem
- The County scale seems to be the partition the furthest from a random distribution of mean T°
- It was shown that different scales can appear as relevant depending on the tackled topic
- The Relative Scalar Deviation being generalized

Josselin *et al.*, 2023, Uncertainties related to real estate price estimation scales, in *Geographic Data Imperfection 2* (Eds.: Batton-Hubert & Pinet) ISTE Wiley

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Thank you for your attention

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ONE SAMPLE

	Sick due to pesticide	Not sick	TOTAL
Orsini Viper	200	800	1000
Apollon Butterfly	50	950	1000
TOTAL	250	1750	2000

Probability to be sick for vipers: 200/1000 = 0,20 = 20 %

Probability to be sick for butterflies: 50/1000 = 0,05 = 5 %

Relative Risk = 0,20/0,05 = 4

(4 times more for vipers)











TWO SEPARATED SAMPLES

Sample 1	Sick	OK	TOTAL
Viper	193	224	417
Butterly	39	45	84
TOTAL	232	269	501

Relative Risk = (193/417) / (39/84) = 1

Sample 2	Sick	OK	TOTAL
Viper	7	576	583
Butterfly	11	905	916
TOTAL	18	1481	1499

Relative Risk = (7/583) / (11/916) = 1











TWO SEPARATED SAMPLES

Sample 1	Sick	OK	TOTAL
Viper	193	224	417
Butterly	39	45	84
TOTAL	232	269	501

Relative Risk = (193/417) / (39/84) = 1

Sample 2	Sick	OK	TOTAL
Viper	7	576	583
Butterfly	11	905	916
TOTAL	18	1481	1499

(same risk)

Relative Risk = (7/583) / (11/916) = 1







Aggregated Relative Risk = 1 ←





≠!

TWO SEPARATED SAMPLES

	Sick due to pesticide	Not sick	TOTAL
Orsini Viper	200	800	1000
Apollon Butterfly	50	950	1000
TOTAL	250	<i>17</i> 50	2000
ONE SAMPLE			

Aggregated Relative Risk = 1 +

(same risk)

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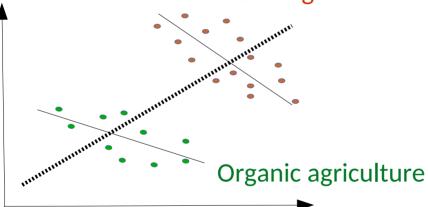


(Spatial) aggregation bias

2 parcels of vine described by V1 et V2

V2 : cost / ha

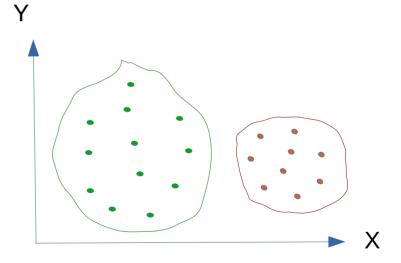
Intensive agriculture



V1: production / ha

In a point plot

Located measures on a map



In geographical space









