# Rituals of Reason: Experimental Evidence on the Social Acceptability of Lotteries in Allocation Problems* 

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#### Abstract

We study collective preferences towards the use of random procedures in allocation mechanisms. We report the results of two experiments in which subjects choose a procedure to allocate a reward to half of them. The first possibility is an explicitly random device: the result of a lottery. The second is an equally unpredictable procedure with identical rate of success, but not involving any explicit randomization. We identify an aversion to lotteries, in particular against procedures that are reminiscent of meritocratic ones. In line with the literature, we also find evidence of a control premium in most procedures.


Keywords: lotteries, mechanism design, allocation problems, procedures, tie-breaking rule

JEL-Code: D01, D78, D91

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## 1 Introduction

A frequent issue in allocation problems is the choice of a procedure to break ties when no better criterion is available. While many of the solutions provided by economists involve a role for randomization and the use of lotteries, ${ }^{1}$ random tie-breakers are often unpopular in practice. In USA law, the Administrative Procedure Act (APA) forbids the use of random devices in administrative decisions. ${ }^{2}$ Several European countries have witnessed a large opposition to school choice mechanisms involving randomization. ${ }^{3}$ The UK government's "school admission code (2014)" explicitly states (p.14) that "local authorities must not use random allocation as the principal oversubscription criterion". ${ }^{4}$ Medical researchers often report difficulties to run random control trials (RCT), as doctors oppose allocating a treatment at random. They only see randomization as ethical if a state of equipoise is reached: a consensus that none of the possible outcomes of the lottery is ex ante better than the others (Lilford and Jackson, 1995). It is however very difficult for a doctor not to seek an additional rationalization instead of declaring two possible treatments to be ex ante equivalent (Donovan et al., 2014). Even economists are reluctant to use lotteries for themselves. To allocate funding, academics spend a lot of time ranking research projects of similar quality and reach conclusions that are as good as random (Cole et al., 1981; Graves et al., 2011; Pier et al., 2018). ${ }^{5}$

The goal of this paper is to understand the social acceptability of lotteries in allocation problems. In two incentivized experiments using a neutral framing, we ask subjects to choose between two unpredictable procedures, one clearly random and the other not, to allocate a reward within their group. We document an aversion to the explicit use of lotteries, in particular when the alternative follows the rituals of reason (Elster, 1989, p.37): ${ }^{6}$ albeit unpredictable, it is reminiscent of meritocratic procedures and subjects may be able to interpret it as such. We also replicate the result that subjects have a

[^1]preference for control over most procedures (Owens et al., 2014; Bartling et al., 2014; Bobadilla-Suarez et al., 2017; Ferreira et al., 2020). This concept is related to the idea of illusion of control (Langer, 1975; Sloof and von Siemens, 2017), which explains this preference by overconfidence. By running treatments with and without control on all procedures, we however show that control alone does not explain the aversion to lotteries we identify.

Our subjects win a prize if they rank among the top half in the procedure chosen by their group. Experiment 1 is pre-registered and uses a representative sample of the US population. ${ }^{7}$ We run two treatments, each with a different alternative non-random procedure. The explicitly random procedure uses the results of a public lottery in Washington DC. The first non-random procedure is a variant of a rock-paper-scissors ( $R P S$ ) game, designed to be unpredictable and devoid of psychological thrill: we ask subjects to give us in advance a list of five moves, that we then play against the moves of all other participants. To ensure no difference in the entertainment value of our procedures, all subjects play RPS regardless of the chosen procedure, and learn the result alongside the lottery one. Despite having to trust that we implement the procedure correctly, $78 \%$ of the subjects have a preference for $R P S$ over the more transparent lottery. In the second non-random procedure, Time, we ask subjects to provide a time of the day, and then implement an intractable and unnecessarily complicated algorithm. Despite being unpredictable, non-transparent and not reminiscent of any existing meritocratic procedure, a significant majority ( $55.9 \%$ ) still chooses Time over the lottery. These preferences are strict: subjects are willing to pay in order to choose the non-random procedure.

In Experiment 2, we use a large online non-representative sample. Subjects still choose between our lottery and a non-explicitly random procedure, and we run all treatments - including the lottery - with and without control. We also use a third non-random procedure reminiscent of meritocratic ones, Paintings, in which subjects have to guess, from pairs of paintings selected by the second author of this paper, the ones preferred by the first author.

The results of our non-representative sample are consistent with the representative one of Experiment 1 for the two treatments used in both, albeit with a smaller preference for non-random treatments. The results for Painting are not significantly different from those of RPS. Perhaps surprisingly, we find that while subjects have a preference for control in most procedures, the largest impact is for the lottery (an 11pp difference in the share of subjects choosing that procedure). This result is particularly policy-relevant as adding control on the lottery is something we are not aware exists in any real-world allocation mechanism.

Our paper focuses on a difficulty with the social acceptability of lotteries in the context of the collective choice of a tie-breaking rule in an allocation procedure. There is widespread evidence however that lotteries are perceived, in general, as fair, and help make unequal outcomes socially acceptable (Bolton et al., 2005; Schmidt and Trautmann, 2019). Individuals also often choose to rely on a randomization device to make their choice when indifferent or indecisive between two alternatives (see Agranov and Ortoleva, 2021 for a recent example). This preference for randomization has for instance been observed in university choice (Dwenger et al., 2018), an allocation problem in which many subjects seem however reluctant for the mechanism itself to use lotteries.

[^2]Lotteries appear easier to implement when they apply to allocation problems where the decision-maker is not a subject of the allocation. For instance in the green card lottery or the random tie-breaking rule of the short-term skilled migration programs H 1B in the USA (Pathak et al., 2020), the participants are, by definition, not US citizens. Their preferences have no direct impact on the social acceptability of the procedure by the public voting for it. In the allocation of social housing, the decision to randomize is mostly made by people unlikely to apply for one. When economists run RCTs, they expect the treatment to benefit a share of the population, and to have no effect or even a negative one on those who were not treated (Aldashev et al., 2017; Deaton and Cartwright, 2018; Heckman, 2020). Whether the subjects of the RCTs would have preferred another allocation mechanism is generally not discussed. In other famous cases were lotteries are implemented, it is not as a final tie-breaking rule. The military draft in the USA offers a large number of exemptions after the lottery numbers are drawn, and those were widely used during the Vietnam war (Bailey and Chyn, 2020). In Rotating and Saving Credit Associations (ROSCAs), the lottery (when there is one) typically happens at the beginning of the process (Anderson and Baland, 2002).

Our results are related to the concept of outcome bias, a tendency to interpret success by merit and effort and ignore the role of luck (Loewenstein and Issacharoff, 1994; Frank, 2016; Brownback and Kuhn, 2019). Hence, subjects may be willing to interpret any device that is not explicitly a lottery as more meritocratic, even if it is in practice completely unpredictable. Our results are also a form of source uncertainty: individuals treat uncertainty differently depending on the mechanism generating it (see, for instance, Heath and Tversky, 1991; Fox and Tversky, 1995; Abdellaoui et al., 2011).

Finally, our work is related to the survey evidence of Keren and Teigen (2010) on preferences for lotteries on determining the outcome of hypothetical decisions involving other people. Oberholzer-Gee et al. (1997) also report an opinion poll suggesting that a market mechanism is the only less acceptable procedure than a lottery for the allocation of a nuclear waste facility.

In Section 2, we present the design of the two experiments. Section 3 outlines the main results. We conclude and develop two main suggestions on how to make random tie-breaking rules socially acceptable in Section 4: to make a larger use of unpredictable but "reasonable" procedures, and to give participants as much control as possible on the chosen one.

## 2 Experimental design

In both experiments, all subjects receive a fixed payment ( $\$ 1$ in Experiment 1 and $\$ 0.80$ in Experiment 2), and half of them receive an additional payment ( $\$ 2.00$ and $\$ 1.60$ ). ${ }^{8}$ We also offer additional smaller incentives for specific questions such as the elicitation of beliefs and of the intensity of preferences. 577 participants completed Experiment 1 and 1,324 Experiment 2. The median time spent in Experiment 1 is 5 minutes and 11 seconds, and 4 minutes and 41 seconds in Experiment 2, for a median hourly payment of $\$ 20.87$ and $\$ 19.71$ respectively. Following the results of Snowberg and Yariv (2021) who found no impact of doubling the rewards for a battery of experimental games, we do not expect our results to depend on their exact level.

[^3]The experiment is composed of three parts. In the first part, we introduce two procedures to allocate the additional payment: one is explicitly random, and the other is not. Subjects also choose their strategies for the procedure(s) in which they have control (if any). In the second part, we ask them to vote on which procedure to use to allocate a reward to half of the members of their group. We focus on preferences for allocation problems in which the decision makers are a subject of the mechanism, but not the only one. We do so in contrast to spectator settings in which subjects make decisions for the others without monetary consequence for themselves, and with individual preferences in which a subject's choice only apply to themselve. In Experiment 2, we implement the choice of the majority. In Experiment 1, we also measure the intensity of preferences. To do so, we ask subjects for their preferred procedure and the smallest amount we should pay them to change their choice. We then implement the choice of a subject chosen at random. ${ }^{9}$ In the third part, we ask them incentivized questions on their beliefs about their ranks in the mechanism, and one incentivized measure of their ambiguity aversion (Experiment 2 only). Finally, we ask non-incentivized demographic and feedback questions. We provide screenshots of the complete instructions in Online Appendix F.

The first experiment in chronological order is Experiment 2. We ran it on Amazon Mechanical Turk (AMT), on June 22, July 7 and August 31, 2021. ${ }^{10}$ In this experiment, we randomly allocate subjects among 12 treatments of on average 110 participants. The number of treatments corresponds to all possible choices between one of our three nonrandom procedures and the random one, with or without control. We then pre-registered Experiment 1 and ran it on August 14, 2023, using Prolific on a representative sample of 600 USA residents. ${ }^{11}$ We ran two treatments of this experiment, one with $R P S$ and one with Time as a non-random procedure (with control), offered alongside the lottery without control. We excluded subjects who did not complete the experiment, in line with our pre-registration. We detail the sample in Table 1. We ran both experiments using oTree (Chen et al., 2016).

On top of testing pre-registered hypothesis, the objective of Experiment 1 was to estimate the level of support among the general US population for the most and least popular non-random procedures. We can then combine the results of Experiment 1 with Experiment 2 to understand the marginal impact of adding or removing control.

For all the treatments, we provide subjects with a personal ID code. We then posted all the strategies and results on a website owned by Lancaster University, to allow subjects to check their strategies, as well as our procedures (see screenshots in Online Appendix F). In Experiment 1, we also sent an email to all subjects with the detail of their payments, as well as their results in the two procedures.

We first describe the explicitly random procedure and then the three non-random ones. Each procedure consists in ranking subjects based on a sequence of five actions.

Our explicitly random procedure, the Lottery, is a bet on whether each of the 5 numbers in the next-day results of a state lottery, the DC-5 lottery, are odd or even.

[^4]Table 1: Size of the sample for each treatment.

| Non-random procedure | Control? | Experiment 1 Experiment 2 |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Lottery |  |  |
|  |  | No | No | Yes |
| RPS | No | - | 152 | 140 |
|  | Yes | 291 | 98 | 99 |
| Paintings | No | - | 123 | 114 |
|  | Yes | - | 97 | 89 |
| Time | No | - | 106 | 91 |
|  | Yes | 286 | 102 | 113 |

Note: When yes for control, subjects chose the sequence used in the corresponding procedure.

As the lottery is run by a third party, our subjects should in theory trust it as being the most transparent we offer. It also satisfies the five properties identified by Eliaz and Rubinstein (2014) as characterizing a fair random procedure. In the treatment without control, we draw a sequence for them, as public authorities do in actual school choice lotteries. ${ }^{12}$

The following three procedures RPS, Paintings and Time do not exhibit explicit randomness in the allocation. We however argue (and show in Section 3.2) that they are in practice unpredictable. The objective of RPS and Paintings is to respect Elster (1989)'s rituals of reason: subjects make decisions that can be interpreted as meaningful and are easy to understand, although they have no predictable influence on the outcome. In contrast, Time is designed to appear arbitrary and meaningless.

In $R P S$, each action is taken from the set \{Rock, Paper, Scissors\}. In the treatment without control, we provide the subjects with a sequence of actions. In the treatment with control, we let them choose their actions. We play the actions of each subject against all other subjects in their treatment. As in the traditional game, Rock wins against Scissors, Scissors against Paper, and Paper against Rock. A subject wins if they win more rounds than their opponent. ${ }^{13}$ We then rank all participants by their number of wins. The main difference with a traditional RPS game is that all strategies are chosen in advance and a subject uses the same strategies against all other subjects. The objective is to ensure people do not choose RPS for its entertainment value: regardless of the chosen procedure, they get to "play" RPS in the same way.

In Paintings, we explain to the subjects that one of the two experimenters, Renaud, has chosen 5 pairs of paintings, each pair by the same artist, and that the second experimenter, Elias, has chosen in each pair his favourite painting. To win, subjects have to guess the paintings chosen by Elias. We rank participants according to the number of

[^5]Table 2: The influence of control on the percentage of participants choosing the non-random procedure.

|  | Experiment 1 |  | Experiment 2 |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Control in |  |  |  |  |  |
| $\quad$ Non-Random | Yes | Yes | Yes | No | No |
| $\quad$ Lottery | No | No | Yes | No | Yes |
| Non-random procedure |  |  |  |  |  |
| RPS | $78.0 \%$ | $74.5 \%$ | $58.6 \%$ | $65.1 \%$ | $51.4 \%$ |
| Paintings | - | $66.0 \%$ | $59.6 \%$ | $62.6 \%$ | $51.8 \%$ |
| Time | $55.9 \%$ | $46.1 \%$ | $41.6 \%$ | $48.1 \%$ | $37.4 \%$ |

Note: Share over all subjects in the sample.
paintings they have guessed correctly. The first half gets the reward. In case of tie, we use the first pair of paintings, then the second and so on. In the treatment without control, we provide the subjects with a sequence of actions. In the treatment with control, we let them choose their actions. Subjects could download a password protected PDF copy of Elias' choices and we revealed the password alongside the results of the experiments.

We designed the third non-random procedure, Time, in a way to make it look as unreasonable as possible: an arbitrary, complicated, and intractable alogrithm. In the treatment without control, we provide subjects with a code corresponding to the last five digits of a time, in hours, minutes and seconds. In the treatment with control, we asked subjects to choose their time, which is then transformed into their five digits code. We rank the code of all subjects using a long list of rules loosely related the number of odd digits. The full algorithm is given in Online Appendix E.

## 3 Results

### 3.1 Choice of Procedure

We report the share of subjects choosing the non-random procedure in each treatment in Table 2. First, we look at the two treatments of the pre-registered Experiment 1, based on a representative sample of the US population. When $R P S$ is offered with control, and Lottery without, $78 \%$ choose the former. This serves as a baseline of the collective preference for a non-random procedure following the rituals of reason. When Time is offered, $56 \%$ choose it over the lottery. In both treatments, we can reject the hypothesis that this share is equal to $50 \%$ (with $p<0.001$, using a one sample two-sided t-test). As predicted, $R P S$ is chosen by a majority of subjects and the share of subjects choosing $R P S$ over the lottery is significantly higher than the share choosing Time. ${ }^{14}$ The fact that Time is chosen by a majority of subjects however goes against our predictions, showing stronger than expected aversion to lotteries.

The results are not statistically different when removing the small number of subjects ( $3.1 \%$ of our sample) who revealed being indifferent between the two procedures by asking for a minimum amount of $\$ 0$ to change their choice. We detail the intensity of preferences

[^6]and run robustness checks without the indifferent subjects in the Online Appendix A.
As expected from the literature, preferences for the different procedures change with control. The two first columns of Experiment 2 in Table 2 correspond to the treatments in which subjects have control over the non-random procedure. The second and fourth columns correspond to the treatments in which they have control over the lottery. The results show that control indeed matters for the lottery, as on average subjects choose the lottery with control $50 \%$ of the time, compared to $39 \%$ without in Experiment 2. ${ }^{15}$ It is not the case in general for the non-random procedures, with a p-value of the Fisher test of 0.27 .

When we separate the procedures, control does not matter for Time (p-value of 0.92) but matters for the procedures following the rituals of reason with an increase in choice of these by 7 pp ( p -value of the Fisher test of 0.046). This is nonetheless lower than the 11pp increase in the choice of the Lottery. A possible explanation for the higher preference for control in lotteries than in non-random procedures is that subjects may misunderstand the probability of some sequences in a lottery, such as five Even for instance. We however show in the Online Appendix D that the actual sequence has no influence on the choice of procedure in the treatments without control over the lottery.

The results of Experiment 2 therefore suggest that roughly half of the social preference for $R P S$ observed in Experiment 1 - the difference between the share of the subjects choosing that procedure and $50 \%$ - can be explained by control, and the other half by following the rituals of reason. While keeping the same treatment conditions on control, preference for RPS over Lottery is roughly similar in Experiment 2 and 1 (the p-value of the Fisher test is 0.48 ). Similarly, the share of subjects choosing Time is higher in Experiment 1, but not significantly so (p-value of the Fisher test is 0.083 ). We should thus interpret the results of Experiment 2 as providing an upper bound for the social acceptability of lotteries in the USA.

### 3.2 All our procedures are unpredictable

To verify our conjecture that the results of all our procedures are unpredictable, we compare our incentivized measures of beliefs with the actual performance of subjects in Table 3. We find no significant correlation. Each row corresponds to a different treatment, denoted by the name of the corresponding non-random procedure.

For $R P S$ in particular, we were concerned that some subjects could exploit a wellknown bias of RPS: too many players choose Rock ( $37.3 \%$ in our sample), and not enough Scissors (27.9\%), in particular in the first round ( $44.5 \%$ and $19.8 \%$ respectively), so that Paper (chosen by $34.8 \%$ of subjects) gives a higher probability of winning. ${ }^{16}$ While we cannot rule out that some participants correctly guessed they were going to win, on aggregate their predictions are as good as random.

[^7]Table 3: Proportion of subject believing that they will win or lose in a procedure who win in that procedure.

|  | Non-Random |  |  |  |  | Lottery |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :---: | :---: |
| Treatment | Win | Lose | P-value $^{1}$ |  | Win | Lose | P-value $^{1}$ |  |  |
| Experiment | $\mathbf{1}$ |  |  |  |  |  |  |  |  |
| $\quad$ RPS | $41.2 \%$ | $53.6 \%$ | 0.067 |  | $52.3 \%$ | $49.1 \%$ | 0.637 |  |  |
| $\quad$ Time | $49.1 \%$ | $52.8 \%$ | 0.624 |  | $54.1 \%$ | $47.1 \%$ | 0.239 |  |  |
| Experiment | $\mathbf{2}$ |  |  |  |  |  |  |  |  |
| $\quad$ RPS | $49.5 \%$ | $50.5 \%$ | 0.853 |  | $48 \%$ | $51.5 \%$ | 0.462 |  |  |
| $\quad$ Paintings | $52.7 \%$ | $49.5 \%$ | 0.541 |  | $52.6 \%$ | $51.2 \%$ | 0.843 |  |  |
| $\quad$ Time | $53.5 \%$ | $48.8 \%$ | 0.369 |  | $48.9 \%$ | $52.2 \%$ | 0.552 |  |  |

Note: We have a similar table for subjects losing in the procedure.
${ }^{1} \mathrm{P}$-value of the Fisher test of the proportion of subjects saying they will win or lose in a procedure predicting if they win or lose.

### 3.3 Beliefs and overconfidence

We now look at the role of beliefs and overconfidence in determining individual choices. In both experiments, what matters is a subject's expectation of being in the first half of the participants of their treatment group. In Experiment 1, where subjects have control over non-random procedures only, we observe some "illusion of control:" $72.5 \%$ of our subjects expect to win RPS, versus $56.0 \%$ the Lottery in one treatment, and $62.9 \%$ expect to win Time versus $53.5 \%$ the Lottery in the other. ${ }^{17}$ In Experiment 2 however, where all procedures are offered with and without control, we see no significant difference between the expectations of success RPS and the Lottery and Time and the Lottery, respectively. ${ }^{18}$

To explore further to what extent choices are driven by preferences or beliefs, Table 4 reports the difference between the choices of subjects who expect to win in the nonrandom procedure but not in the lottery, and, conversely, of those who only expect to win the lottery. The numbers for Experiment 2 aggregate the different levels of control.

In both experiments, we see that subjects act to a certain extent according to their beliefs, as they are significantly more likely to pick $R P S$ if they expect to win in that procedure only, than if they expect to only win in the Lottery. This is not the case for Time, a procedure for which it is arguably more difficult to convince yourself of your ability to win.

Beliefs are however far from explaining everything. In Experiment 1, we see for instance that a significant majority of subjects $(66.7 \%)$ prefer $R P S$ even when they expect to only win in the Lottery. This preference is a strict one: they are willing to pay in order to choose a procedure in which they nonetheless think they are less likely to win. ${ }^{19}$

[^8]Table 4: Percentage of subjects who choose the non-random procedure conditional on their beliefs of winning only in one procedure but not the other.

|  | Subjects expect to win only in |  |  |
| :--- | :--- | :--- | :--- |
| Treatment | Non-random | Lottery | P-value $^{1}$ |
| Experiment | $\mathbf{1}$ |  |  |
| RPS | $84.9 \%$ | $63.2 \%$ | 0.010 |
| Time | $63.8 \%$ | $45.2 \%$ | 0.117 |
| Experiment 2 |  |  |  |
| RPS | $79.1 \%$ | $46.2 \%$ | $<0.001$ |
| Paintings | $67 \%$ | $43.7 \%$ | 0.003 |
| Time | $50.6 \%$ | $45.3 \%$ | 0.530 |

Note: The first column represents subjects expecting to win in the non-random procedure but not in the lottery, the second column the reverse.
${ }^{1} \mathrm{P}$-value of the Fisher test of the proportions of subjects who believe they will win in one procedure but not the other who choose the non-random procedure being equal.

In the Online Appendix C, we confirm all these results in a regression analysis. We also provide additional information on the subjects' demographics and relative preferences, as well as more detail on our sample characteristics in Online Appendix B.

## 4 Discussion and Conclusion

We start this section by providing two recommendations on the implementation of random tie breakers based on the results of our experiment and the previous literature. They are aimed at any designer willing to implement a random (explicitly or not) tie-breaking rule. We then conclude by discussing limitations of the paper and avenues for future research.

Our first recommendation is that mechanism designers looking for a random tiebreaking rule should aim first for unpredictable procedures that are not explicitly random, but look reasonable enough to be more socially acceptable. In the case of university admission, tie-breakers are often needed between students with similar high school grades. In the United States, as in the United Kingdom and the Republic of Ireland, the high school marking standard indeed places students in broad categories, usually in letter form. One of the reasons for such a standard is that it allows conveying information about student achievement without putting the burden on markers to provide an impossible level of precision (Schneider and Hutt, 2014). While providing more precision may be seen as unfair or arbitrary in the context of high school results, it may actually prove useful to avoid requiring an additional tie-breaking rule for the mechanisms using these results. In other words, if the difference between a $787 / 1000$ and a $788 / 1000$ is not meaningful and does not tell us anything about the relative abilities of two students, it nonetheless provides a non-random procedure following the rituals of reason on which students had some control.

Simple tweaks can also be used to reduce the probability of ties without increasing the burden put on markers. In the allocation mechanism of high school students to university
in the Republic of Ireland, students are allocated a random number to break ties among those who have an identical CAO score, a metric composed of the sum, over 6 subjects, of their leaving certificate, composed of letter grades translated into marks out of $100 .{ }^{20}$ Following complaints that random tie-breaking procedures were unfair and too frequent, the Irish government revised that metrics in 2015 with as objective to reduce the number of ties. Before the reform, possible marks on a given subject where all multiple of 5 . After the reform, the possible marks are $\{0,37,46,56,66,77,88,100\}$, so that summing them up leads to much fewer ties. The government does not hide that those numbers have been chosen with the main objective of reducing ties, and have no particular meaning. ${ }^{21}$

Our second recommendation is that, when the only choice is to have a lottery, it may be worth trying to give participants some control on it. In line with Eliaz and Rubinstein (2014) subjects need to understand what is happening, the procedure must be familiar. As shown with our arbitrary procedure Time, subjects do not prefer an obscure algorithm over a clear lottery once they have control on the latter. For instance, in France, the algorithm allocating students to university between 2009 and 2017 was simply not public. Citizens could find out it was a modified deferred acceptance mechanism but were not provided any information on the details. In 2017, the French public authority overseeing data privacy law (CNIL) summoned the French government to explain to participants to the mechanism the precise algorithm used to rank them (decision 201053, 30 August 2017). ${ }^{22}$ The current procedure features a transparent central algorithm, but the algorithms used by universities to rank students are not disclosed. Simply put, citizens have a general idea of what the mechanism aims to achieve and are aware that the choice is made by algorithms, but know little more than that.

It would also help to let people take an active part to the process. The Washington DC "school lottery" is a prime example of a mechanism that publicly embraces the concept of lottery. It is a deferred acceptance mechanism using an individual randomly generated number as a tie-breaker (Abdulkadiroğlu et al., 2017). However, instead of trying to hide the random tie-breaker, the entire procedure goes by the name of "lottery." How random number are being assigned to each parent is however often strikingly non-transparent. For the DC lottery, as well as for the NYC and Denver deferred acceptance mechanisms, official websites simply inform parents that the algorithm generates a random number, but there is no way to actually check how the number is generated, and certainly no involvement of the subjects in the random procedure. We are not aware of any real world mechanism making the random tie-breaking rule transparent and offering some form of control to participants. Our results however suggest that doing so may increase the social acceptability of such lotteries.

We see two main limitations to our approach. The first is inherent to every incentivized study aiming at understanding major choices in life: the experimental stakes we can offer will never be sufficiently high to mimic the real life incentives. We cannot think of any experimental reward one could reasonably offer that would approach the importance for a parent of putting their child in their preferred school, or for a researcher to get a major grant. We however believe that our incentivized approach offers a step in the right direction, allowing to better understand individual choices than what surveys on stated preferences do. The second is that our data on cultural dimensions between subjects

[^9]is limited, and we would certainly learn from running a similar experiment in different countries.

Finally, while making our subjects vote on the choice of procedures aims at replicating democratic processes, we could learn more by adding an important feature of actual democracies: debate and consensus building. Observing how smaller groups of subjects manage to agree on a procedure would clearly improve our understanding of collective preferences.

## References

Abdellaoui, M., Baillon, A., Placido, L., Wakker, P.P., 2011. The rich domain of uncertainty: Source functions and their experimental implementation. American Economic Review 101, 695-723.

Abdulkadiroğlu, A., Agarwal, N., Pathak, P.A., 2017. The welfare effects of coordinated assignment: Evidence from the new york city high school match. American Economic Review 107, 3635-89.

Abdulkadiroglu, A., Angrist, J.D., Narita, Y., Pathak, P.A., 2019. Breaking ties: Regression discontinuity design meets market design .

Agranov, M., Ortoleva, P., 2021. Ranges of randomization. mimeo .
Aldashev, G., Kirchsteiger, G., Sebald, A., 2017. Assignment procedure biases in randomised policy experiments. The Economic Journal 127, 873-895.

Anderson, S., Baland, J.M., 2002. The economics of roscas and intrahousehold resource allocation. The Quarterly Journal of Economics 117, 963-995.

Avin, S., 2019. Mavericks and lotteries. Studies in History and Philosophy of Science Part A 76, 13-23.

Bailey, M.J., Chyn, E., 2020. The demographic effects of dodging the vietnam draft, in: AEA Papers and Proceedings, American Economic Association 2014 Broadway, Suite 305, Nashville, TN 37203. pp. 220-225.

Barnett, A.G., 2016. Funding by lottery: political problems and research opportunities. MBio 7, e01369-16.

Bartling, B., Fehr, E., Herz, H., 2014. The intrinsic value of decision rights. Econometrica 82, 2005-2039.

Basteck, C., Klaus, B., Kübler, D., 2021. How lotteries in school choice help to level the playing field. Games and Economic Behavior .

Bobadilla-Suarez, S., Sunstein, C.R., Sharot, T., 2017. The intrinsic value of choice: The propensity to under-delegate in the face of potential gains and losses. Journal of risk and uncertainty 54, 187-202.

Bolton, G.E., Brandts, J., Ockenfels, A., 2005. Fair procedures: Evidence from games involving lotteries. The Economic Journal 115, 1054-1076.

Brezis, E.S., 2007. Focal randomisation: An optimal mechanism for the evaluation of r\&d projects. Science and Public Policy 34, 691-698.

Brownback, A., Kuhn, M.A., 2019. Understanding outcome bias. Games and Economic Behavior 117, 342-360.

Budish, E., Che, Y.K., Kojima, F., Milgrom, P., 2013. Designing random allocation mechanisms: Theory and applications. American economic review 103, 585-623.

Chen, D.L., Schonger, M., Wickens, C., 2016. otree - an open-source platform for laboratory, online, and field experiments. Journal of Behavioral and Experimental Finance 9, 88 - 97. doi:10.1016/j.jbef.2015.12.001.

Cole, S., Simon, G.A., et al., 1981. Chance and consensus in peer review. Science 214, 881-886.

Deaton, A., Cartwright, N., 2018. Understanding and misunderstanding randomized controlled trials. Social Science \& Medicine 210, 2-21.

Donovan, J.L., de Salis, I., Toerien, M., Paramasivan, S., Hamdy, F.C., Blazeby, J.M., 2014. The intellectual challenges and emotional consequences of equipoise contributed to the fragility of recruitment in six randomized controlled trials. Journal of clinical epidemiology 67, 912-920.

Dwenger, N., Kübler, D., Weizsäcker, G., 2018. Flipping a coin: Evidence from university applications. Journal of Public Economics 167, 240-250.

Eliaz, K., Rubinstein, A., 2014. On the fairness of random procedures. Economics Letters 123, 168-170.

Elster, J., 1989. Solomonic judgements: Studies in the limitation of rationality. Cambridge University Press.

Erdil, A., Ergin, H., 2008. What's the matter with tie-breaking? improving efficiency in school choice. American Economic Review 98, 669-89.

Fang, F.C., Casadevall, A., 2016. Research funding: The case for a modified lottery. mBio 7, e00422-16.

Ferreira, J.V., Hanaki, N., Tarroux, B., 2020. On the roots of the intrinsic value of decision rights: Experimental evidence. Games and Economic Behavior 119, 110-122.

Fox, C.R., Tversky, A., 1995. Ambiguity aversion and comparative ignorance 110, 585603. doi:10.2307/2946693. publisher: Oxford Academic.

Frank, R.H., 2016. Success and luck: Good fortune and the myth of meritocracy. Princeton University Press.

Graves, N., Barnett, A.G., Clarke, P., 2011. Funding grant proposals for scientific research: retrospective analysis of scores by members of grant review panel. Bmj 343, d4797.

Greenberg, D.S., 1998. Chance and grants. The Lancet 351, 686.

Heath, C., Tversky, A., 1991. Preference and belief: Ambiguity and competence in choice under uncertainty 4, 5-28. doi:10.1007/BF00057884.

Heckman, J.J., 2020. Randomization and Social Policy Evaluation Revisited. Technical Report.

Keren, G., Teigen, K.H., 2010. Decisions by coin toss: Inappropriate but fair. Judgment and Decision Making 5, 83.

Kesten, O., Ünver, M.U., 2015. A theory of school-choice lotteries. Theoretical Economics 10, 543-595.

Langer, E.J., 1975. The illusion of control. Journal of personality and social psychology 32, 311.

Lilford, R.J., Jackson, J., 1995. Equipoise and the ethics of randomization. Journal of the Royal Society of Medicine 88, 552.

Loewenstein, G., Issacharoff, S., 1994. Source dependence in the valuation of objects. Journal of Behavioral Decision Making 7, 157-168.

Oberholzer-Gee, F., Bohnet, I., Frey, B.S., 1997. Fairness and competence in democratic decisions. Public Choice 91, 89-105.

Owens, D., Grossman, Z., Fackler, R., 2014. The control premium: A preference for payoff autonomy. American Economic Journal: Microeconomics 6, 138-61.

Pathak, P.A., Rees-Jones, A., Sönmez, T., 2020. Immigration Lottery Design: Engineered and Coincidental Consequences of H-1B Reforms. Technical Report. National Bureau of Economic Research.

Pier, E.L., Brauer, M., Filut, A., Kaatz, A., Raclaw, J., Nathan, M.J., Ford, C.E., Carnes, M., 2018. Low agreement among reviewers evaluating the same nih grant applications. Proceedings of the National Academy of Sciences 115, 2952-2957.

Roumbanis, L., 2019. Peer review or lottery? a critical analysis of two different forms of decision-making mechanisms for allocation of research grants. Science, Technology, \& Human Values 44, 994-1019.

Schmidt, R.J., Trautmann, S.T., 2019. Implementing (un) fair procedures? favoritism and process fairness when inequality is inevitable.

Schneider, J., Hutt, E., 2014. Making the grade: A history of the a-f marking scheme. Journal of Curriculum Studies 46, 201-224.

Sloof, R., von Siemens, F.A., 2017. Illusion of control and the pursuit of authority. Experimental Economics 20, 556-573.

Snowberg, E., Yariv, L., 2021. Testing the waters: Behavior across participant pools. American Economic Review 111, 687-719.

Vermeule, A., 2015. Rationally arbitrary decisions in administrative law. The Journal of Legal Studies 44, S475-S507.

## Appendix

## A Intensity of Preferences in Experiment 1

We only elicited our subjects' willingness to pay to change their choice in Experiment 1. We estimate that a subject is indifferent if they state a $\$ 0$ minimal value for changing their choices.

Table 5: Number of subjects and strength of the preference.

| Non-random |  | Procedure |  | Preference |  |
| :--- | :--- | :--- | :---: | :---: | :---: |
| Procedure |  | Chosen |  | Indifferent | Strict |
| RPS | Lottery |  | 4 | 60 |  |
| RPS | RPS |  | 10 | 217 |  |
| Time | Lottery |  | 3 | 123 |  |
| Time | Time |  | 1 | 159 |  |

Note:
A preference is considered strict if the WTA is non-null.


Figure 1: Minimum amount we would need to pay subjects for them to change their choice. The reward for ranking among the first half of the chosen procedure is $\$ 2$.

Table 5 shows that a very large majority of subjects have strict preferences. These preferences are quite strong, as implied by the distribution of the minimal values asked for in Figure 1. The proportion of the non-random procedure being chosen are not significantly different with or without indifferent subjects. Fisher exact tests of equality of proportion gives a P-value of 1 in the case of $R P S$ and of 0.933 in the case of Time. Table 6 shows that the proportions are always significantly different from $50 \%$, showing that there is a majority of subjects favoring one procedure over the other (in general the non-random one).

Table 6: Proportion of the non-random procedure being chosen in each treatment, when excluding indifferent subjects.

| Procedure $\quad$ Non-Random Chosen | P-value $^{1}$ |  |
| :--- | :---: | :---: |
| Experiment 1 |  |  |
| RPS | $78 \%$ | $<0.001$ |
| Time | $55.9 \%$ | 0.044 |
| Experiment 2 |  |  |
| RPS | $61.8 \%$ | $<0.001$ |
| Paintings | $59.8 \%$ | $<0.001$ |
| Time | $43.4 \%$ | 0.008 |

${ }^{1} \mathrm{P}$-value of the one sample two-sided t-test of equality with $50 \%$.
Table 7: Proportion of subjects of each gender in both experiments.

| Gender | Experiment 1 | Experiment 2 |
| :--- | :---: | :---: |
| Female | $52.5 \%$ | $38.8 \%$ |
| Male | $47.5 \%$ | $61.2 \%$ |

## B Demographics

Table 7 shows the gender repartition in Experiment 1 and 2. There are a lot more males in Experiment 2, whereas Experiment 1 is representative in terms of gender, as expected. Table 8 shows the repartition by age group in Experiment 1 and 2. As Experiment 1 is made on a representative sample of the USA population, it shows that Experiment 2 is far from representative in terms of age. Table 9 shows the repartition of ethnicity using Prolific stratification strategy.

We report the self-declared country of residence of the subjects in Experiment 1 and 2 in Table 10. The majority of our subjects report coming from the United States (79.3\%) in Experiment 2. The second highest country of residence is India (12.2\%).

Table 8: Proportion of subjects in each age group for each experiment.

| Age Group | Experiment 1 | Experiment 2 |
| :--- | :---: | :---: |
| $<25$ | $13.9 \%$ | $3.5 \%$ |
| $25-40$ | $28.9 \%$ | $64.6 \%$ |
| $40-55$ | $25.8 \%$ | $23.2 \%$ |
| $>55$ | $31.4 \%$ | $8.2 \%$ |
| Prefer not to say | - | $0.5 \%$ |
| Note: |  |  |
| Age group used are those elicited in Experiment 2. |  |  |

Table 9: Declared ethnicity in Experiment 1

| Declared Race | Count |
| :--- | :---: |
| Asian | 36 |
| Black | 76 |
| Mixed | 12 |
| Other | 6 |
| White | 447 |

Table 10: Number of subjects from each country in both experiments.

| Country | Experiment 1 | Experiment $2^{1}$ |
| :--- | :---: | :---: |
| Asian | - | $0.2 \%$ |
| Brazil | - | $4.2 \%$ |
| Bulgaria | - | $0.1 \%$ |
| Canada | - | $0.5 \%$ |
| Columbia | - | $0.1 \%$ |
| France | - | $0.3 \%$ |
| Germany | - | $0.3 \%$ |
| India | - | $12.2 \%$ |
| Italy | - | $1.5 \%$ |
| Portugal | - | $0.1 \%$ |
| Spain | - | $0.2 \%$ |
| Sweden | - | $0.1 \%$ |
| The Netherlands | - | $0.1 \%$ |
| Turkey | - | $0.1 \%$ |
| UAE | - | $0.2 \%$ |
| USA | $100 \%$ | $79.3 \%$ |
| Ukraine | - | $0.1 \%$ |
| United Kingdom | - | $0.5 \%$ |

${ }^{1}$ In Experiment 2, the country is residence is selfdeclared. We have tried to reconstruct the country from their declaration, but it was sometimes difficult.

## C Regression analysis

We perform a linear regression analysis of the choice of the non-random procedure in both experiments in Table 11. The coefficients represents variations (in percentage points) of the share of subjects who chose a non-random procedure. The baseline is given with $R P S$ with control and Lottery without control. We find that the procedures following the rituals of reason are indistinguishable from each other, whereas the arbitrary one (Time) is less chosen ( 22 percentage points in Experiment 1). Control matters more for the lottery than for the non-random procedures. It increases the choice of the lottery by around 10 percentage point, and the absence of control over the non-random procedures by 5 percentage points. For some subjects, the belief that they will win in one of the procedure and not the other drives their choices. It represents a fraction of around 20 percent of the sample. There is however no correlation between the expected relative ranking in the measures and the choice of procedure. This last result is not unexpected, as the only payoff-relevant ranking is whether or not a subject is in the first half of the participants, not the relative ranking.

## D Influence of the Given Sequence in the Lottery

It is possible that participants are influenced in their choices by the strategies we choose for them when they have no control over a procedure. For instance, they may believe that the sequence Even, Even, Even, Even, Even or Odd, Odd, Odd, Odd, Odd is less likely to happen than the sequence Even, Odd, Even, Odd, Even. This incorrect understanding of probabilities is more likely to happen when facing the Lottery, as there is no particular order for the Paintings or for the Time procedure. In $R P S$, subjects would have to be sophisticated in their strategies to form a belief about it, and it does not appear to be the case.

To investigate this issue, we restrict ourselves to the sample of subjects without control on the lotteries, which is all participants of Experiment 1 and 678 participants in Experiment 2. We first count the number of Even in the sequence we gave them. Table 12 shows the proportions of subjects choosing each non-random procedure depending on the number of Even in their sequence. The sample sizes vary widely: subjects are much more likely to have a sequence with 3 Even than 5 (even if any given sequence is equally likely).

First we find no difference between the treatment of Odd and Even. There is no significant difference between the share of subjects choosing the non-random procedure when awarded a sequence of 0 or 5 Even. The same holds between 1 and 4 or 2 and 3 , respectively, according to a Fisher test of equal proportions. So in Table 12, Odd and Even are treated the same by participants, which is reassuring.

We can therefore group up sequences of 0 and 5 Even together, as well as sequences of 1 and 4 and 2 and 3 . It yields to Table 13. The proportions of each non-random procedure being chosen are never significantly different between 0 and 1 or 2 Even in the sequence, according to a Fisher test of equal proportions.

Finally, we run a regression without demographic controls restricted to subjects without control on the lottery. We add dummy for getting 0 or 1 even (or odd) and use as a baseline getting 2 or 3 evens. The results given by Table 14 shows that the given sequence does not significantly influence choices. The significant variables also do not change, and neither do their magnitudes, compared to the coefficients in Table 11 without

Table 11: Regression of the choice of the non-random procedure.

|  | Non-random Procedure Chosen |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Experiment 1 |  | Experiment 2 |  |
| (Intercept) | $\begin{gathered} \hline 0.780 \\ (0.032) \end{gathered}$ | $\begin{gathered} \hline 0.848 \\ (0.070) \end{gathered}$ | $\begin{gathered} \hline 0.695 \\ (0.032) \end{gathered}$ | $\begin{gathered} \hline 0.699 \\ (0.078) \end{gathered}$ |
| Time | $\begin{gathered} -0.221^{* * *} \\ (0.038) \end{gathered}$ | $\begin{gathered} -0.223^{* * *} \\ (0.038) \end{gathered}$ | $\begin{gathered} -0.190^{* * *} \\ (0.033) \end{gathered}$ | $\begin{gathered} -0.191^{* * *} \\ (0.033) \end{gathered}$ |
| Paintings |  |  | $\begin{aligned} & -0.028 \\ & (0.032) \end{aligned}$ | $\begin{aligned} & -0.024 \\ & (0.032) \end{aligned}$ |
| Win in Non-Random ${ }^{\text {a }}$ | $\begin{aligned} & 0.077+ \\ & (0.045) \end{aligned}$ | $\begin{aligned} & 0.076+ \\ & (0.046) \end{aligned}$ | $\begin{gathered} 0.114 * * * \\ (0.033) \end{gathered}$ | $\begin{gathered} 0.112^{* * *} \\ (0.033) \end{gathered}$ |
| Win in Lottery ${ }^{\text {a }}$ | $\begin{aligned} & -0.126^{*} \\ & (0.060) \end{aligned}$ | $\begin{aligned} & -0.132^{*} \\ & (0.060) \end{aligned}$ | $\begin{gathered} -0.099^{* *} \\ (0.035) \end{gathered}$ | $\begin{gathered} -0.098^{* *} \\ (0.035) \end{gathered}$ |
| Indifferent | $\begin{aligned} & -0.127 \\ & (0.109) \end{aligned}$ | $\begin{aligned} & -0.120 \\ & (0.110) \end{aligned}$ |  |  |
| No Control on Non-Random |  |  | $\begin{gathered} -0.053^{*} \\ (0.027) \end{gathered}$ | $\begin{array}{r} -0.050+ \\ (0.027) \end{array}$ |
| Control on Lottery |  |  | $\begin{gathered} -0.103^{* * *} \\ (0.027) \end{gathered}$ | $\begin{gathered} -0.107^{* * *} \\ (0.027) \end{gathered}$ |
| Male |  | $\begin{aligned} & -0.044 \\ & (0.038) \end{aligned}$ |  | $\begin{aligned} & 0.052+ \\ & (0.027) \end{aligned}$ |
| Demographic Controls ${ }^{\text {b }}$ | No | Yes | No | Yes |
| Num.Obs. | 577 | 577 | 1324 | 1317 |
| R2 | 0.073 | 0.079 | 0.059 | 0.065 |
| R2 Adj. | 0.066 | 0.064 | 0.055 | 0.058 |

$+\mathrm{p}<0.1,{ }^{*} \mathrm{p}<0.05,{ }^{* *} \mathrm{p}<0.01,{ }^{* * *} \mathrm{p}<0.001$
Note:
No demographic control is significant.
${ }^{\text {a }}$ Win in XX is a dummy for when a subjects believe they win the XX procedure and NOT in the alternative one.
${ }^{\mathrm{b}}$ Demographic controls: Age categories, Being White (in Experiment 1)

Table 12: Choices of the non-random procedure when subjects have no control over the lottery, by the number of Even in the sequence given to them.

|  | Number of evens |  |  |  |  |  | P -values |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 0 | 1 | 2 | 3 | 4 | 5 | $0 \sim 5$ | 1~4 | 2~3 |
| Experiment 2 |  |  |  |  |  |  |  |  |  |
| RPS | 40.0\% | 63.6\% | 65.0\% | 77.5\% | 76.9\% | 50.0\% | $>0.999$ | 0.234 | 0.108 |
| Paintings | 66.7\% | 63.4\% | 70.0\% | 62.2\% | $52.2 \%$ | 75.0\% | $>0.999$ | 0.433 | 0.375 |
| Time | 28.6\% | 39.3\% | 49.2\% | 54.3\% | $34.4 \%$ | 62.5\% | 0.315 | 0.791 | 0.604 |
| Experiment 1 |  |  |  |  |  |  |  |  |  |
| RPS | 88.9\% | 86.5\% | 78.6\% | 78.3\% | 71.9\% | 60.0\% | 0.505 | 0.139 | >0.999 |
| Time | 37.5\% | 53.2\% | 60.4\% | 51.7\% | 56.5\% | 85.7\% | 0.119 | 0.836 | 0.290 |
| Aggregate | 52.5\% | 61.9\% | 65.1\% | 64.9\% | 61.3\% | 67.6\% | 0.237 | 0.918 | $>0.999$ |
| Note: |  |  |  |  |  |  |  |  |  |

Table 13: Choices of the non-random procedure when subjects have no control over the lottery, by the number of Even/odd in the sequence given to them.

|  | Number of evens/odds |  |  |  | P-values |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Non-random procedure | 0 | 1 | 2 |  | $0 \sim 1$ | $0 \sim 2$ | $1 \sim 2$ |  |
| Experiment 2 |  |  |  |  |  |  |  |  |
| $\quad$ RPS | $43.8 \%$ | $69.9 \%$ | $70.9 \%$ |  | 0.081 | 0.045 | 0.882 |  |
| $\quad$ Paintings | $71.4 \%$ | $59.4 \%$ | $65.5 \%$ |  | 0.548 | 0.773 | 0.436 |  |
| $\quad$ Time | $46.7 \%$ | $36.7 \%$ | $51.9 \%$ |  | 0.558 | 0.789 | 0.062 |  |
| Experiment 1 |  |  |  |  |  |  |  |  |
| $\quad$ RPS | $78.6 \%$ | $77.2 \%$ | $78.4 \%$ |  | $>0.999$ | $>0.999$ | 0.881 |  |
| $\quad$ Time | $60.0 \%$ | $54.8 \%$ | $56.2 \%$ |  | 0.785 | $>0.999$ | 0.898 |  |
| Aggregate | $59.5 \%$ | $61.6 \%$ | $65.0 \%$ |  | 0.795 | 0.374 | 0.251 |  |

Note:
P-values of the Fisher exact test of equal proportion for each procedure between each number of sequence.

Table 14: Choice of the non-random procedure, adding controls for the sequences.

demographic controls. It means that despite some anecdotal evidence, how the sequence looks may only be marginally taken into account by our participants in their choices of procedure.

## E The Algorithm for Time

We show here the version with control of the arbitrary algorithm used in our Time procedure.

We will ask you to choose a time (in 24 hours format). For each player we will record the five last digits, that we denote as your "code". For instance, 10 hours, 26 minutes and 31 seconds become " 02631 ".

We have developed an algorithm ranking all participants based on their code (We expect around 100 participants today). Among others, it takes into account whether your code is above or below the median. You can read the details of the algorithm by
clicking on this button.
The details below are hidden by default, but can be revealed by clicking on a button. We use the following algorithm to rank the codes:

1. We will count for all the players in the experiment the number $n$ of odd digits of the code, with 0 counting as even. In the example, the number of odd digits is $\mathrm{n}=$ 2.
2. We will then rank everyone according to the number $n$ (a higher $n$ yield a higher rank). We call this rank your "code rank".
3. For the tied players with the same number $n$ of odd digits, we rank them by the statistical frequency of the first digit of the "code". We will give a higher rank to those with the lowest frequency, then to the second lowest one, until there are no more number left (tied frequencies are bundled together). If there is still a tie, we repeat the same procedure with the second digit of the code. And so till the last digit.
4. We then determine the winner as follows:

- CASE 1: If strictly more players have $\mathrm{n}<2.5$ than $\mathrm{n}>2.5$ (i.e., if the median is below 2.5): your award rank is the same as your code rank. All the $50 \%$ higher ranked players in the award rank win the reward.
- CASE 2: If strictly fewer players have $\mathrm{n}<2.5$ than $\mathrm{n}>2.5$ (i.e., if the median is above 2.5): your award rank is the revert of the code rank. If there are N participants and your code rank was j , then your award rank is $\mathrm{N}+1-\mathrm{j}$. All the $50 \%$ higher ranked players in the award rank win the reward.

5. In the unlikely event that a tie remains at the end of the procedure, exactly at the $50 \%$ mark, all tied players will receive a reward.

## F Screenshots (for Online Publication)

## F. 1 Experiment

In this section, we show the screenshots of the different pages of Experiment 1. Each screen appeared once, except that only one of Time or RPS was shown. Beliefs were elicited for Time in the same way as for RPS. The experiment was programme using oTree (Chen et al., 2016). The screens were very similar for Experiment 2, with two important differences. First, we did not ask for the intensity of their preferences. Second, we had demographic questions that we do not need in this experiment, as they are provided by Prolific.

## Welcome!

Thank you for participating in this experiment! You will receive a payment of at least $\$ 1.00$ for your participation. You are in a group of 100 participants for this experiment.

The experiment is divided into three stages. Your choices in each of these stages may increase your total earnings.

1. In the first stage, you will choose five actions in a Rock-Paper-Scissors game (RPS hereafter).
2. In the second stage, you will choose between two options that may give you a reward of $\$ 2.00$. Your chances of getting the reward will depend on your choices in stage 1.
3. In the third stage, we will ask you some questions regarding your previous choices. By answering these questions, you may earn $\$ 0.20$.

For transparency, once we have computed the results and anonymized the choices, we will send you a personal code and a website link through Prolific.

Please note that once you have clicked the next button on a page, you won't be able to go back and change your answers. In other words, once validated, your answers are definitive.

## The First Mechanism: Rock, Paper, Scissors

## Your Actions

## Rules

Rock, Paper, Scissors (RPS) is a game originating from ancient China. Its rules are simple. You have to choose one of three actions: Rock, Paper or Scissors. Your opponent also has to choose one of these three actions. Then:

- Rock smashes Scissors;
- Scissors cuts Paper;
- Paper covers Rock;
- If both of you choose the same action, it is a tie.

We ask you to provide us with a sequence of 5 actions of RPS. At the end of the day, we will play your actions against all 99 other members of your group. The game will be such that each time your first action is pitted against your opponent's first action, your second action against their second action, and so on. The rules for winning are simple: each time your action beats the opponent's action, you gain 1 point. If your action is beaten, however, you lose 1 point. In case of a tie, nothing happens.

The points are then summed up, you win if:

- You have a positive number of points, which means you have won more often than your opponent;
- Your number of point is null (i.e., you and your opponent have won the same number of rounds) and you have won the tie breaking rule. 1

We will then rank all participants according to the number of wins. In case of tie, we use the result of the games among tied players only.

For the first round, which action do you want to choose?
Rock
Paper
Scissors

For the second round, which action do you want to choose?
Rock
Paper
Scissors

For the third round, which action do you want to choose?

- Rock

Paper
Scissors

For the fourth round, which action do you want to choose?
Rock
Paper
Scissors

For the fifth round, which action do you want to choose?
Rock

- Paper

Scissors

1: The winner of the tie-breaking rule is the first winner of a round.

2: In the unlikely event where tied players have chosen exactly the same strategy, we use a tie-breaking procedure based on your time of arrival. You can find the details here.

## The Second Mechanism: Time

## Rules

We will ask you to choose a time (in 24 hours format). For each player we will record the five last digits, that we denote as your "code". For instance, 10 hours, 26 minutes and 31 seconds become " 02631 ".

We have developed an algorithm ranking all 100 players of your group based on their code. Among others, it takes into account whether your code is above or below the median. You can read the details of the algorithm by clicking on this button.

## Details

## Your Choices

Please choose your hour:
$\square$

Please choose your minutes:
$\square$

Please choose your seconds:
$\square$

## The Second Mechanism: DC-5 Lottery

## Your Code

Our algorithm has generated the following code for you: Odd, Odd, Even, Even, Odd.

## Rules

It is a bet over whether each of the 5 numbers of the day after tomorrow's draw of the DC-5 Lottery is either Odd or Even (the draw of the November 17, 2023 at 12.50PM, in Washington DC time zone).

For each number, you get one point if your code is correct. We will rank all 100 members of your group by the number of points they got.

If you are among the first half, you win. If you are among the second half, you lose. 1

## Choosing the Mechanism

## The next two pages

One of the next two questions will be used to determine which mechanism determines the allocation of the reward to the 100 players of your group.

We have attributed unique personal numbers between 0 and 99 to all players in your group. Your number is 71 . If your number is drawn the day after tomorrow by the mid-day DC-2 lottery (the draw of the November 17, 2023 at 12.50PM, in Washington DC time zone), we will use your chosen mechanism to determine the rewards in your group. Otherwise, we will use the mechanism chosen by the player whose number was drawn by the lottery. If the first number of the day after tomorrow's mid-day DC-3 lottery is below 5 (if it is $0,1,2,3$ or 4 ), then we will use this question to determine the mechanism. Otherwise, we will use the next question.

You can now choose whether you prefer to get a chance to win the reward of $\$ 2.00$ according to:

- The results from your RPS sequence of actions.
- The results of your 5 digits code against tomorrow's DC-5 Lottery.

In both cases, you win if you are among the first half of all participants in your group.

Which mechanism do you choose?
Rock, Paper, Scissors
DC-5 Lottery

## Changing your choice?

In the previous page, you chose the Rock, Paper, Scissors to allocate the rewards of $\$ 2.00$ to the participants in your group. In this page, you may earn an additionnal amount of up to $\$ 2.00$.

What is the minimal amount we should pay you to change your choice?

We have built the payments so that it is in your best interest to tell us truthfully what is the minimal amount we should pay you to change your choice. The details are accessible when clicking on the link below.

## Details

If the first number of the day after tomorrow's mid-day DC-3 lottery is 5 or above (if it is $5,6,7,8$ or 9 ), then we will use this question to determine the mechanism. Otherwise, we will use the previous question. If this question is used and your number is drawn in the day after tomorrow's DC-2 mid-day lottery, here is how it works.

First, we use the number of the day after tomorrow's mid-day DC-3 lottery, which is comprised between 0 and 999 and divide it by 200. The rest of this division is the threshold amount (in cents).

- If the threshold amount is lower than the amount you gave us above, we will use Rock, Paper, Scissors (the choice you made in the previous page).
- If the threshold amount is higher than the amount you gave us, we pay you the threshold amount, and we will use the DC-5 Lottery.


## Winning in Rock, Paper, Scissors

One of the next three questions will be randomly selected for payment. ${ }^{1}$
If this question is selected for payment, you can earn an additional $\$ 0.20$ if your guess is correct (and $\$ 0.00$ otherwise). ?
You have the sequence of actions in Rock, Paper, Scissors: Rock, Paper, Scissors, Rock, Rock. We will rank your performance according to the algorithm described here. We will play your sequence against all other members of your group and rank all the participants according to the sequence.

Do you think you will be ranked among the:

| Bottom half | Top half |
| :---: | :---: | Next

1: To determine the question used for payment, we add one to the rest of the division of the number drawn by the DC-4 Lottery of the day after tomorrow at mid-day (the draw of November 17, 2023 at 12.50PM).

2: If you are ranked exactly in the middle, you will earn the reward no matter which option you choose.

## Winning in DC-5 Lottery

If this question is selected for payment, you can earn an additional $\$ 0.20$ if your guess is correct (and $\$ 0.00$ otherwise). $\mathbf{2}$
Your sequence of guesses for the DC-5 lottery is Odd, Odd, Even, Even, Odd. We will rank its performance in guessing the DC-5 lottery results against all other codes of members of your group, according to the algorithm described here.

Do you think you will be ranked among the:
Bottom half $\quad$ Top half

## Next

2: If you are ranked exactly in the middle, you will earn the reward no matter which option you choose.

## Winning the Reward

If this question is selected for payment, you can earn an additional $\$ 0.20$ if you guess in which mechanism you are ranked higher (and $\$ 0.00$ otherwise). 1
We have previously described to you two mechanisms, the DC-5 Lottery (details here) and Rock, Paper, Scissors (details here).

Do you think your rank will be higher in?
DC-5 Lottery
Rock, Paper, Scissors

## Understanding your choices

We would like to better understand your choices, the following questions have no right or wrong answer.
Why did you choose the Rock, Paper, Scissors to get your reward?
I think I have more chances of winning in this mechanism.
I prefer this mechanism, but for other reasons.

If you wish, you may explain in more details your choice.

| Do you agree or disagree with the following <br> statements.Strongly <br> disagree |
| :--- |
| Rock, Paper, Scissors is more meritocratic than <br> DC-5 Lottery. <br> Rock, Paper, Scissors is fairer than DC-5 Lottery. <br> Reither agree nor <br> disagreeAgree Strongly <br> agree |
| Rock, Paper, Scissors is easier to understand <br> than DC-5 Lottery. |

## Next

## Thank You!

You have completed this experiment, thank you for participating!

Your reward will be paid on your Prolific account in the next 4 days, after the lotteries are drawn and the mechanisms played accordingly. The anonymized results will be available on the website https://eliasbcd.github.io/experiment/experiment1.html.

Your personal code is 9 h8assid. We will send you the detailed payments and personal code through an email to your Prolific account. To tell us (and Prolific) that you have completed the experiment, please click here.

## F. 2 Results

We publicly showed the anonymized results of the experiment. Here is a screenshot of one result page.

## Results from Experiment 1

You can find here the results from Experiment 1 carried on by $\quad$ and
The different draws of the DC-Lottery of August 16 at 12.50pm are:

- DC-2: 4,9 . The participant number selected is 49 ;
- DC-3: 3,2 , 8 , the rest in the division by 200 is 128 , which is the threshold amount (in cents). The first number is below 5 , so the question used is the first (and the threshold amount does not matter).
- DC-4: $9,6,6,6$, so the rest in the division by 3 is 0 and the question for predictions selected is the first one.
- DC-5: 7, 8, 8, 5, 4, so the sequence to match is Odd, Even, Even, Odd, Even.

Due to a mistake in the implementation of the experiment, you are in group of approximately 50 participants, instead of the 100 originally intended. In each group, the mechanism chosen to compute the reward may be different.

The selected mechanism to attribute the reward in Group 1 is Rock, Paper, Scissors. The choices of all members of Group 1 are below.

| About You |  | Payment | About the Mechanisms |  |  |  |  |  | Predictions |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Code | Number |  | Lottery Sequence | RPS <br> Sequence | Your Choice | Your Threshold | Lottery <br> Ranks | RPS Ranks | RPS | Lottery | Mechanism Preferred |
| gpyun1b2 | 70 | \$3.20 | Even, Even, Even, Even, Odd | Paper, <br> Paper, <br> Paper, <br> Paper, <br> Paper | Rock, <br> Paper, Scissors | \$0.41 | 30 | 1 | Top half | Bottom half | Rock, Paper, Scissors |
| fnl7ptzq | 37 | \$3.00 | Even, Odd, <br> Odd, Odd, Odd | Paper, <br> Scissors, Paper, Rock, Rock | Rock, <br> Paper, Scissors | \$1.00 | 45 | 2 | Top half | Top half | Rock, Paper, Scissors |
| 1 frnbpdj | 6 | \$3.20 | Odd, Even, Even, Odd, Even | Paper, <br> Rock, Scissors, Rock, Paper | Rock, <br> Paper, Scissors | \$1.00 | 1 | 3 | Top half | Top half | DC-5 <br> Lottery |


[^0]:    *We thank Micael Castanheira, Konstantinos Georgalos, Rustamdjam Hakimov, Kim Kaivanto, Dorothea Kubler, Patrick Legros, Jean-Marc Tallon and Eyal Winter for useful discussions and comments. We thank seminar participants at Lancaster University, ESA 2020 world meeting, University Paris 8, University Paris Dauphine, University of Kent, and Université libre de Bruxelles for feedback and suggestions. We declare no competing interests.
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[^1]:    ${ }^{1}$ Randomization is widely accepted by economists as the fairest way to deal with ties in allocation problems, albeit there is a debate on the best way to implement it (see Budish et al., 2013 for a review, and Erdil and Ergin, 2008; Kesten and Ünver, 2015; Abdulkadiroglu et al., 2019 for the case of school choice). Lotteries can also have efficiency benefits: Basteck et al. (2021) show for instance that the introduction of a lottery quota strengthens truth-telling in the deferred acceptance mechanism.
    ${ }^{2}$ The APA [requires court to] set aside agency action that is "arbitrary, capricious, [or] an abuse of discretion." (Vermeule, 2015, p.475) The US Supreme court (Judulang v. Holder, 132 S. Ct. 476, 485 [2011] [Kagan, J.]) rules that the use of a random device is arbitrary and thus illegal (p.11), regardless of how costly it is to provide a rationalization (p.21).
    ${ }^{3}$ In France, the system of allocation of students to universities (APB) used lotteries to break ties when capacity was reached. It was criticized and replaced in 2018 by another mechanism with a lot more criteria. In French-speaking Belgium, a 2009 attempt to randomly allocate students in oversubscribed high schools lasted only a year after being dubbed by some parents and the media the "lottery law" and replaced by a set of criteria still in use today.
    ${ }^{4}$ This debate followed the introduction in 2008 of a random tie-breaking rule in Brighton and Hove, after a criterion of distance to the school. Lotteries are also used as last-resort criteria in cities such as Birmingham. The issue remains a political talking point, and the 2017 Conservative Party manifesto explicitly committed the government to (p.50): "never introduce a mandatory lottery-based school admissions policy."
    ${ }^{5}$ Proposals to save time and money by adding random tie breakers to the criteria used for peer review (Greenberg, 1998; Brezis, 2007; Fang and Casadevall, 2016; Roumbanis, 2019; Avin, 2019) are largely ignored (Barnett, 2016). Notable exceptions are limited to small or exploratory grants, such as the Health Research Council of New Zealand "Explorer" grants, The Volkswagen fundation in Germany and the small grants from the British Academy.
    ${ }^{6}$ "We have a strong reluctance to admit uncertainty and indeterminacy in human affairs. Rather than accept the limits of reason, we prefer the rituals of reason."

[^2]:    ${ }^{7}$ Both our experiments have received IRB approval from Lancaster University and the code and data will be made available in a public repository before publication. Our pre-registration is available at https://aspredicted.org/3yy5p.pdf

[^3]:    ${ }^{8}$ The higher nominal amounts in Experiment 1 (August 2023) than in Experiment 2 (July 2021) reflect the inflation during the period.

[^4]:    ${ }^{9}$ We incentivize truthful revelation of the intensity of preferences in the following way. Denote by $T$ the smallest amount a subject reports to be willing to accept to change her choice. If the subject is selected, we draw a random number $D \in(0, \$ 2)$. If $D<T$ we keep the initial choice of the subject. If $D \geq T$ we transfer an amount $D$ to the subject and implement the other procedure. We could not use the standard voting procedure of Experiment 1, because measuring the intensity of preferences would then imply measuring the perceived probability of being pivotal.
    ${ }^{10}$ We also ran a pilot experiment on April 29 and May 5, 2020, detailed in a previous version of this work, Lancaster Economics Working Papers Series, vol. 2022/005.
    ${ }^{11}$ The pre-analysis plan report is available in a separate online appendix.

[^5]:    ${ }^{12}$ For instance, the official website https://www.schools.nyc.gov reports that a real example of random number generated by the NYC school lottery is B51920AF-F1C6-40EC-8E9A-3E1E50CB13BB. A student or parent can log into their account and see the number, but has no influence on it.
    ${ }^{13}$ In case of a tie, the first winner of a round wins the game. In the rare event where both players have chosen exactly the same strategy, we consider it as neither a win nor a loss.

[^6]:    ${ }^{14}$ The p-value of the Fisher test of the proportion of subjects choosing the non-random procedure in both treatments being different is $<0.001$.

[^7]:    ${ }^{15}$ The p-value of the Fisher Exact test of equality of the proportion is $<0.001$.
    ${ }^{16}$ The bias is similar to the choices of players on the website https://roshambo.me (We thank Lasse Hassing for giving us access to their data). They choose Rock $35.3 \%$ of the time in the first round, and Scissors $29.1 \%$ of the time. The difference in magnitude is likely to arise because players on the website chose to play a game of RPS, contrary to our subjects, and play actual games instead of giving a set of five choices in advance. They are therefore more likely to play regularly and have a notion of the best strategies, at least for some of them.

[^8]:    ${ }^{17}$ The difference are significant according to a Fisher exact test of equal proportions, with a p-value $<0.001$ in the first case and of 0.027 in the second.
    ${ }^{18} \mathrm{P}$-value of the same Fisher tests are 0.74 and 0.53 respectively.
    ${ }^{19}$ As we offered subjects a binary choice of a bet on whether they expect to win or lose in a given procedure, we cannot rule out that some subjects expect to win in both procedures with an identical probability. However, if they reveal a strictly positive willingness to pay to keep their chosen procedure, we can conclude that they have a strict preference for one procedure over the other that is not explained by differences in expectations of winning.

[^9]:    ${ }^{20}$ Central Application Office, "Random number - How it works" on cao.ie, retrieved August, 2, 2021.
    ${ }^{21}$ Central Application Office, The New Common Point Scale, on cao.ie, retrieved August, 2, 2021.
    ${ }^{22}$ See also the description of the algorithm of Admission Post Bac (APB) provided by matching in practice, https://www.matching-in-practice.eu/university-admission-practices-france/.

