

# Estimation of multivariate generalized gamma convolutions, application to dependence structure modelling in insurance.

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Estimation of multivariate generalized gamma convolutions, application to dependence structure modelling in insurance.

Véronique Maume-Deschamps, joint work with Oskar Laverny, Esterina Masiello, Didier Rullière. Partially funded by SCOR







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A (re-)insurer observes many (positive) random variables that represent losses from several portfolios of risks. A major topic is the evaluation and modeling of the dependency between those risks. Indeed, the properties of this dependency has a huge impact on the behavior of the total loss, and therefore on the capital management of the company.

Risks can be heavy-tailed, skewed, and are usually dependant with each other, including in the extremes.

The infinite divisibility may be a wanted property.

### Why infinite divisibility ?

#### Definition (ID r.v. X)

- X is *n*-divisible if and only if there exists  $X_1, ..., X_n$  i.i.d. such that  $X = X_1 + ... + X_n$
- **2** X is *ID* iff this can be done for all n.
- **3**  $X_i$  is called a *piece* of X.
  - Dependencies are largely unobservables in reinsurance;
  - Experts are providing information as divisibility parameters: "20% of this is driven by the same risk factors as 10% of that.." this means comonotonic pieces accross marginals.

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### Company's Solvency



SC (Solvency Capital): regulatory capital calculated using a risk measure that the re-insurer must have to absorb potential losses. For example, in the Solvency II directive of the European Union, it is the VaR(99.5%) of the P&L distribution.

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Risk factors mod	delisation		

In the internal modeling point of view, the distributions of several losses  $X_1, \ldots, X_d$  are supposed to be known, but the dependence structure between them must still be evaluated and taken into account to assess the variability and behavior of the total loss:  $\sum_{i=1}^{d} X_i.$ 

It might be achieved by a risk factors decomposition:  $X_i = \sum_{i=1}^{n} X_{i,j}$ ,

where  $X_{i,j}, X_{\ell,j}$  are either independent or comonotonic. Divisibility allows to interpret the  $X_{i,j}$  as pieces of  $X_i$ .  $\implies$  multivariate generalized Gamma convolutions give an answer to that question.

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### Recall Gamma distributions

### Well known class of infinite divisible distributions.

 $X \sim \mathcal{G}(k, \theta)$  if its density function is:



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### Multivariate Gamma convolutions classes $\mathcal{G}_{d,n}$ and $\mathcal{G}_d$

$$\begin{split} \mathbf{x}, \mathbf{t} &\in \mathbb{R}^{d}, \, \mathbf{Recall:} \, \operatorname{Cumulant} \, \operatorname{generating} \, \operatorname{function:} \\ \mathcal{K}(\mathbf{t}) &= \ln \left( \mathbb{E} \left( e^{\langle \mathbf{t}, \mathbf{X} \rangle} \right) \right). \\ \mathbf{Multivariate} \, \operatorname{Thorin} \, \operatorname{Classes} \, ^{1} \\ \mathbf{X} &\sim \mathcal{G}_{d,n}(\alpha, \mathbf{s}) \Leftrightarrow \mathcal{K}(\mathbf{t}) = -\sum_{i=1}^{n} \alpha_{i} \ln \left( 1 - \langle \mathbf{s}_{i}, \mathbf{t} \rangle \right) \\ \mathbf{X} &\sim \mathcal{G}_{d}(\nu) \Leftrightarrow \mathcal{K}(\mathbf{t}) = -\int \ln \left( 1 - \langle \mathbf{s}, \mathbf{t} \rangle \right) \nu(\partial \mathbf{s}). \end{split}$$

#### Example

All gammas, log-Normals, Paretos,  $\alpha$ -stables, and sums and/or products of these are in  $G_1$ .

#### No estimation procedures were available before our work.

<sup>&</sup>lt;sup>1</sup> Lennart Bondesson (Nov. 2009). en. In: Journal of Statistical Planning and Inference

### Motivation: interpretability of $\mathcal{G}_{d,n}$ models

 $X \in \mathcal{G}_{d,n}$  has an additive risk-factor structure: there exists gamma random variables  $G_1, ..., G_n$  such that:

$$\begin{pmatrix} X_1 \\ \dots \\ X_d \end{pmatrix} = \begin{pmatrix} s_{1,1} & \dots & \dots & s_{1,n} \\ \dots & \dots & \dots & \dots \\ s_{d,1} & \dots & \dots & s_{d,n} \end{pmatrix} \cdot \begin{pmatrix} G_1 \\ \dots \\ G_n \end{pmatrix}$$
 For all  $i \in 1, \dots, n$ ,  

$$\begin{matrix} \mathbf{G}_i \sim \mathcal{G}_{1,1}(\alpha_i, 1) \\ \mathbf{G}_i \sim \mathcal{G}_1, \dots, \mathcal{G}_n \text{ are independent.} \end{matrix}$$

**Goal:** Estimate  $\nu$  from observations of the random vector **X**. **Pb:** It is a deconvolution problem, which is numerically hard (can also be interpreted as a solution-less moment problem for the Thorin measure, see MFK <sup>2</sup>).

<sup>&</sup>lt;sup>2</sup> Justin Miles, Edward Furman, and Alexey Kuznetsov (2019). In: Variance

### The orthonormal Laguerre basis of $L_2(\mathbb{R}^d_+)$

Consider the following orthonormal basis of  $L_2(\mathbb{R}^d_+)$  called Laguerre basis <sup>3</sup>: let  $\mathbf{p} \in \mathbb{N}^d$ ,  $\phi_{\mathbf{p}}(\mathbf{x}) = \prod_{i=1}^d \phi_{p_i}(x_i)$ , with  $\phi_p(x) = \sqrt{2} \sum_{k=0}^p {p \choose k} \frac{(-2x)^k}{k!} e^{-x}$ . Any square integrable distribution density f can be expended as :

$$f(\mathbf{x}) = \sum_{\mathbf{p} \in \mathbb{N}^d} a_{\mathbf{p}} \phi_{\mathbf{p}}(\mathbf{x})$$
 where  $a_{\mathbf{p}} = \int \phi_{\mathbf{p}}(\mathbf{x}) f(\mathbf{x}) \partial \mathbf{x}$ 

<sup>3</sup> Florian Dussap (2021). In: Journal of Statistical Planning and Inference

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### Estimation

Consider a random sample  $X_1, \ldots, X_N$  of  $X \in \mathcal{G}_d(\alpha, \mathbf{s})$ . Let

$$\widehat{\mathbf{a}_{\mathbf{p}}} = \frac{1}{N} \sum_{i=1}^{N} \phi_{\mathbf{p}}(\mathbf{X}_{i}).$$

Consider the integrated square error loss:  $L(\alpha, \mathbf{s}) = \sum_{\mathbf{k} \le \mathbf{m}} (\widehat{a_{\mathbf{k}}} - a_{\mathbf{k}}(\alpha, \mathbf{s}))^2 \text{ and}$ 

 $(\widehat{\alpha}, \widehat{\mathbf{s}}) = \operatorname{argmin}_{(\alpha, \mathbf{s})} L(\alpha, \mathbf{s}).$ 

Remark that the theoretical Laguerre coefficients may be computed using the formula:

$$a_{\mathbf{k}} = \langle \phi_{\mathbf{k}}, f \rangle = \mathbb{E}\left(\phi_{\mathbf{k}}(\mathbf{X})\right) = \sqrt{2}^{d} \sum_{\mathbf{k} \leq \mathbf{p}} {\binom{\mathbf{p}}{\mathbf{k}}} \frac{(-2)^{|\mathbf{k}|}}{\mathbf{k}!} \mathbb{E}\left(X^{\mathbf{k}} e^{-|\mathbf{X}|}\right).$$

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Consistency			

We have the following consistency result <sup>4</sup> under the thechnical *well-behaved* asumption (which holds for univariate Gamma convolutions, independent multivariate Gamma convolution and invertible linear transformations of them e.g.).

#### Theorem

If **X** is drawn from an  $\epsilon$ -well-behaved density  $f_{\alpha,\mathbf{s}} \in \mathcal{G}_d(\alpha,\mathbf{s})$  then,

$$\|f_{(\widehat{\alpha},\widehat{\mathbf{s}})} - f_{\alpha,\mathbf{s}}\|_2^2 \xrightarrow[N \to \infty]{n \to \infty} 0.$$

<sup>4</sup> Oskar Laverny et al. (Jan. 2021). In: *Electronic Journal of Statistics* 

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Dimensionality is	ssues		

The loss  $L(\alpha, \mathbf{s})$  is too costly to work with when *d* gets large. Indeed, the number of coefficients to compute is given by

$$D(m,d) = \sum_{i=1}^m \binom{i+d-1}{d-1}.$$

which is exponentially increasing in d and therefore unusable in high dimension (ok until d = 4 or 5.

Improvement of the method by Oskar Laverny <sup>5</sup> by using an approximation of an integrated loss function and random projections (tested for d = 1000).

<sup>5</sup> Oskar Laverny (2022). In: arXiv preprint arXiv:2203.13741

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### Multivariate log-Normal



Figure: MLN case: n = 10

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### Multivariate log-Normal



Figure: MLN case: n = 20

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### Tail dependency: A Clayton copula



Figure: Clayton case: n = 20

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### Conclusions

- The univariate Thorin class is wide: LN, Pareto, α-stable, (some) Weibull, ...
- The Multivariate analogue provides an asymmetrical dependence structure which include tail dependency, and can take a lot of different shapes.
- Deconvolution is a hard inverse problem, and estimation of these distributions is complicated.
- The final additive risk-factor model gives easy interpretation of parameters and easy aggregation schemes. Should be shortly used by SCOR in some parts of their internal model.
- Improvement to go to high dimension in Oskar' work <sup>6</sup> and the open-source Julia package <sup>7</sup> provides the implementation.

<sup>&</sup>lt;sup>6</sup> Oskar Laverny (2022). In: *arXiv preprint arXiv:2203.13741* <sup>7</sup> Oskar Laverny (Mar. 2021). Version v0.1







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### Thank you

# Thanks for your attention.