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# Empirical Risk Minimization with $f$ -Divergence Regularization in Statistical Learning

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H. Vincent Poor

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**Abstract:** This report presents the solution to the empirical risk minimization with  $f$ -divergence regularization, under mild conditions on  $f$ . Under such conditions, the optimal measure is shown to be unique and to always exist. The solution is presented as a closed-form expression of the Radon-Nikodym derivative of the optimal probability measure with respect to the reference measure. Examples for particular choices of the function  $f$  are presented. For some choices, existing results are obtained as special cases of the main result. These include the unique solutions to the empirical risk minimization with relative entropy regularization (Type-I and Type-II).

**Key-words:** Empirical Risk Minimization;  $f$ -divergence; Regularization; Inductive Bias; Statistical Learning.

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# Minimisation du Risque Empirique avec Régularisation par $f$ -Divergences dans l'Apprentissage Statistique

**Résumé :** Ce rapport présente la solution au problème de minimisation du risque empirique avec régularisation par une  $f$ -divergence dans des faibles contraintes sur  $f$ . Dans ces conditions, la mesure optimale s'avère unique et existe toujours. La solution est présentée comme une expression sous forme fermée de la dérivée de Radon-Nikodym de la mesure de probabilité optimale par rapport à la mesure de référence. Des exemples de choix particuliers de la fonction  $f$  sont présentés. Pour certains choix, des résultats existants sont obtenus comme cas particuliers du résultat principal. Il s'agit notamment des solutions uniques aux problèmes de minimisation du risque empirique avec régularisation par entropie relative (Type-I et Type-II).

**Mots-clés :** Minimisation du Risque Empirique,  $f$ -divergence, Régularisation, apprentissage statistique.

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## 1 Introduction

Empirical Risk Minimization (ERM) is a fundamental principle in machine learning. It is a tool for selecting a model from a given set by minimizing the empirical risk, which is the average loss or error induced by such a model on each of the labelled patterns available in the training dataset [1, 2]. In simpler terms, ERM aims to find a model that performs well on a given training dataset. However, ERM is prone to overfitting [3–5], which affects the generalization capability of the resulting optimal model [6–8]. To remediate this phenomenon, the solution of the ERM shall exhibit a small sensitivity to variations in the training dataset, which is often obtained via regularization [9–14].

In statistical learning theory, the ERM problem consists in the minimization of the expected empirical risk over a subset of all probability measures that can be defined upon the set of models. In this case, regularization is often obtained by adding (up to a constant factor) a *statistical distance* from the optimization measure to a *reference measure* to the expected empirical risk (w.r.t. the optimization measure). Such a statistical distance is essentially a non-negative measure of dissimilarity between the optimization measure and the reference measure, which might be a  $\sigma$ -finite measure and not necessarily a probability measure, c.f. [12, 15] and [16]. A key observation is that such a reference measure often represents prior knowledge and/or desired features on the solution.

A typical example of a statistical distance is an  $f$ -divergence. The notion of  $f$ -divergence was introduced in [17] and further studied in [18, 19]. A popular  $f$ -divergence is the relative entropy [20], which due to its asymmetry, leads to two different problems when it is used for regularizing the ERM problem. Those problems are known as Type-I and Type-II ERM with relative entropy regularization (ERM-RER). In the Type-I ERM-RER, the regularizer is the relative entropy of the optimization measure with respect to (w.r.t.) the reference measure [12]. Alternatively, in the Type-II ERM-RER, the regularizer is the relative entropy of the reference measure w.r.t. the optimization measure [13]. The Type-I ERM-RER problem exhibits a unique solution, which is a Gibbs probability measure. In the case in which the reference measure is a probability measure, the existence of the solution is always ensured. Alternatively, in the case in which it is a  $\sigma$ -finite measure, the existence of a solution is subject to strict conditions. The Type-II ERM-RER problem also exhibits a unique solution when the reference measure is a probability measure. Such a solution is a probability measure that exhibits properties that are analogous to those of the Gibbs probability measure, as pointed out in [13]. Nonetheless, despite these similarities, Type-I ERM-RER appears to be the most popular regularized ERM problem in statistical learning theory. See for instance, [15, 16, 21–30] and references therein.

Optimization problems with  $f$ -divergence regularization has been explored in [31, 32] and [33]. Nonetheless, the problem of ERM with  $f$ -divergence regulariza-

tion (ERM- $f$ DR), with a general  $f$  function, has not been studied before. Only the special cases,  $f(x) = -x \log(x)$  (Type-I ERM-RER) and  $f(x) = -\log(x)$  (Type-II ERM-RER), mentioned above have attracted particular attention. This work presents the solution to the ERM- $f$ DR problem using a method of proof that differs from those in [31, 32] and [33] and goes along the lines of the methods in [12, 13] and [14], which rely on the notion of the Gateaux derivative [34].

The remainder of this report is organized as follows. Section 2 presents the standard ERM problem and introduces the notation. The ERM- $f$ DR problem and its solution, which is the main contribution of this work, is presented in Section 3. Section 4 presents examples for specific choices of the function  $f$ . Finally, Section 5 draws the main conclusions and finalizes this report.

## 2 Empirical Risk Minimization Problem

Let  $\mathcal{M}$ ,  $\mathcal{X}$  and  $\mathcal{Y}$ , with  $\mathcal{M} \subseteq \mathbb{R}^d$  and  $d \in \mathbb{N}$ , be sets of *models*, *patterns*, and *labels*, respectively. A pair  $(x, y) \in \mathcal{X} \times \mathcal{Y}$  is referred to as a *labeled pattern* or *data point*. Given the data points  $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$  with  $n \in \mathbb{N}$ , a *dataset* is represented by the tuple  $((x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)) \in (\mathcal{X} \times \mathcal{Y})^n$ .

Let the function  $h : \mathcal{M} \times \mathcal{X} \rightarrow \mathcal{Y}$  be such that the label assigned to a pattern  $x \in \mathcal{X}$  according to the model  $\theta \in \mathcal{M}$  is  $h(\theta, x)$ . Then, given a dataset

$$\mathbf{z} = ((x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)) \in (\mathcal{X} \times \mathcal{Y})^n, \quad (1)$$

the objective is to obtain a model  $\theta \in \mathcal{M}$ , such that, for all  $i \in \{1, 2, \dots, n\}$ , the label assigned to pattern  $x_i$ , which is  $h(\theta, x_i)$ , is “close” to the label  $y_i$  specified by the training dataset. This notion of “closeness” is formalized by the function

$$\ell : \mathcal{Y} \times \mathcal{Y} \rightarrow [0, +\infty), \quad (2)$$

such that the loss or risk induced by choosing the model  $\theta \in \mathcal{M}$  w.r.t. the labelled pattern  $(x_i, y_i)$ , with  $i \in \{1, 2, \dots, n\}$ , is  $\ell(h(\theta, x_i), y_i)$ . The risk function  $\ell$  is assumed to be nonnegative and satisfy  $\ell(y, y) = 0$ , for all  $y \in \mathcal{Y}$ .

The *empirical risk* induced by a model  $\theta$  with respect to the dataset  $\mathbf{z}$  in (1) is determined by the function  $L_{\mathbf{z}} : \mathcal{M} \rightarrow [0, +\infty)$ , which satisfies

$$L_{\mathbf{z}}(\theta) \triangleq \frac{1}{n} \sum_{i=1}^n \ell(h(\theta, x_i), y_i). \quad (3)$$

The ERM problem w.r.t. the dataset  $\mathbf{z}$  in (1) consists of the optimization problem:

$$\min_{\theta \in \mathcal{M}} L_{\mathbf{z}}(\theta). \quad (4)$$

The set of solutions to such a problem is denoted by

$$\mathcal{T}(\mathbf{z}) \triangleq \arg \min_{\theta \in \mathcal{M}} L_{\mathbf{z}}(\theta). \quad (5)$$

Note that if the set  $\mathcal{M}$  is finite, the ERM problem in (4) has a solution, and therefore, it holds that  $|\mathcal{T}(\mathbf{z})| > 0$ . Nevertheless, in general, the ERM problem does not always have a solution. That is, there exist choices of the loss function  $\ell$  and the dataset  $\mathbf{z}$  that yield  $|\mathcal{T}(\mathbf{z})| = 0$ .

### 3 The ERM with $f$ -Divergence Regularization

#### 3.1 Preliminaries

For the ease of notation, the expected empirical risk with respect to a given measure is expressed via the following functional  $R_{\mathbf{z}}$ , defined hereunder.

**Definition 3.1** (Expected Empirical Risk). *The expectation of the empirical risk  $L_{\mathbf{z}}(\boldsymbol{\theta})$  in (3), when  $\boldsymbol{\theta}$  is sampled from a probability measure  $P \in \Delta(\mathcal{M})$ , is determined by the functional  $R_{\mathbf{z}} : \Delta(\mathcal{M}) \rightarrow [0, +\infty)$ , such that*

$$R_{\mathbf{z}}(P) = \int L_{\mathbf{z}}(\boldsymbol{\theta}) dP(\boldsymbol{\theta}). \quad (6)$$

Similarly, a particular notation is used for the  $f$ -divergences, as shown in the following definition.

**Definition 3.2** ( $f$ -divergence [19]). *Let  $f : [0, \infty) \rightarrow \mathbb{R}$  be a convex function with  $f(1) = 0$  and  $f(0) = \lim_{x \rightarrow 0^+} f(x)$ . Let  $P$  and  $Q$  be two probability measures on the same measurable space, with  $P$  absolutely continuous with  $Q$ . The  $f$ -divergence of  $P$  w.r.t.  $Q$ , denoted by  $D_f(P\|Q)$ , is*

$$D_f(P\|Q) \triangleq \int f\left(\frac{dP}{dQ}(\boldsymbol{\theta})\right) dQ(\boldsymbol{\theta}), \quad (7)$$

where the function  $\frac{dP}{dQ}$  is the Radon-Nikodym derivative of  $P$  with respect to  $Q$ .

The set of probability measures that can be defined upon the measurable space  $(\mathcal{M}, \mathcal{B}(\mathcal{M}))$ , with  $\mathcal{B}(\mathcal{M})$  being the Borel  $\sigma$ -field on  $\mathcal{M}$ , is denoted by  $\Delta(\mathcal{M})$ . Given a probability measure  $Q \in \Delta(\mathcal{M})$  the set containing exclusively the probability measures in  $\Delta(\mathcal{M})$  that are absolutely continuous with  $Q$  is denoted by  $\Delta_Q(\mathcal{M})$ . That is,

$$\Delta_Q(\mathcal{M}) \triangleq \{P \in \Delta(\mathcal{M}) : P \ll Q\}, \quad (8)$$

where the notation  $P \ll Q$  stands for the measure  $P$  being absolutely continuous with respect to measure  $Q$ .

#### 3.2 Problem Formulation

The ERM- $f$ DR problem is parametrized by a probability measure  $Q \in \Delta(\mathcal{M})$ , a positive real  $\lambda$ , and an  $f$ -divergence (Definition 3.2). The measure  $Q$  is referred to as the *reference measure*,  $\lambda$  as the *regularization factor*, and  $f$  as the

*regularization function.* Given the dataset  $\mathbf{z} \in (\mathcal{X} \times \mathcal{Y})^n$  in (1), the ERM- $f$ DR problem, with parameters  $Q$ ,  $\lambda$  and  $f$ , consists of the following optimization problem:

$$\min_{P \in \Delta_Q(\mathcal{M})} R_{\mathbf{z}}(P) + \lambda D_f(P \| Q). \quad (9)$$

### 3.3 Solution to the ERM- $f$ DR

The solution of the ERM- $f$ DR problem in (9) is presented in the following theorem under the assumption that the function  $f$  is strictly convex and differentiable.

**Theorem 3.1.** *Let the function  $f : (0, +\infty) \rightarrow \mathbb{R}$  in (9) be strictly convex and differentiable, with  $f(1) = 0$ . Denote by  $\dot{f} : (0, +\infty) \rightarrow \mathcal{A}$  and  $\dot{f}^{-1} : \mathcal{A} \rightarrow (0, +\infty)$ , with  $\mathcal{A} \subseteq \mathbb{R}$ , the derivative of  $f$  and the inverse of the function  $\dot{f}$ , respectively. The optimization problem in (9) possesses a unique solution, denoted by  $P_{\Theta | \mathbf{Z}=\mathbf{z}}^{(Q, \lambda)}$ , whose Radon-Nikodym derivative with respect to the probability measure  $Q$  satisfies for all  $\boldsymbol{\theta} \in \text{supp } Q$*

$$\frac{dP_{\Theta | \mathbf{Z}=\mathbf{z}}^{(Q, \lambda)}}{dQ}(\boldsymbol{\theta}) = \dot{f}^{-1}\left(-\frac{\beta + L_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}\right), \quad (10)$$

if there exists a  $\beta$  in  $\left\{t \in \mathbb{R} : \forall \boldsymbol{\theta} \in \text{supp } Q, 0 < \dot{f}^{-1}\left(-\frac{t + L_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}\right) < +\infty\right\}$  such that

$$\int \dot{f}^{-1}\left(-\frac{\beta + L_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}\right) dQ(\boldsymbol{\theta}) = 1. \quad (11)$$

*Proof:* The optimization problem in (9) can be re-written in terms of the Radon-Nikodym derivative of the optimization measure  $P$  w.r.t. the measure  $Q$ , which yields:

$$\min_{P \in \Delta_Q(\mathcal{M})} \int L_{\mathbf{z}}(\boldsymbol{\theta}) \frac{dP}{dQ}(\boldsymbol{\theta}) dQ(\boldsymbol{\theta}) + \lambda \int f\left(\frac{dP}{dQ}(\boldsymbol{\theta})\right) dQ(\boldsymbol{\theta}) \quad (12a)$$

$$\text{s.t.} \quad \int \frac{dP}{dQ}(\boldsymbol{\theta}) dQ(\boldsymbol{\theta}) = 1. \quad (12b)$$

The remainder of the proof focuses on the problem in which the optimization is over the function  $\frac{dP}{dQ} : \mathcal{M} \rightarrow [0, \infty)$ , which represents the Radon-Nikodym derivative of  $P$  w.r.t.  $Q$ . Hence, instead of optimizing the measure  $P$ , the optimization is over the function  $\frac{dP}{dQ}$ . This is due to the fact that for all  $P \in \Delta_Q(\mathcal{M})$ , the Radon-Nikodym derivative  $\frac{dP}{dQ}$  is unique up to sets of zero measure w.r.t.  $Q$ . Let  $\mathcal{M}$  be the set of measurable functions  $\mathcal{M} \rightarrow \mathbb{R}$  with respect to the measurable spaces  $(\mathcal{M}, \mathcal{F})$  and  $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$  that are absolutely integrable with respect to  $Q$ . That is, for all  $\hat{g} \in \mathcal{M}$ , it holds that

$$\int |\hat{g}(\boldsymbol{\theta})| dQ(\boldsymbol{\theta}) < \infty. \quad (13)$$

Hence, the optimization problem of interest is:

$$\min_{g \in \mathcal{M}} \int \mathbf{L}_z(\boldsymbol{\theta})g(\boldsymbol{\theta}) dQ(\boldsymbol{\theta}) + \lambda \int f(g(\boldsymbol{\theta}))dQ(\boldsymbol{\theta}) \quad (14a)$$

$$\text{s.t. } \int g(\boldsymbol{\theta})dQ(\boldsymbol{\theta}) = 1. \quad (14b)$$

Let the Lagrangian of the optimization problem in (14) be  $L : \mathcal{M} \times \mathbb{R} \rightarrow \mathbb{R}$  such that

$$L(g, \beta) = \int \mathbf{L}_z(\boldsymbol{\theta})g(\boldsymbol{\theta}) dQ(\boldsymbol{\theta}) + \lambda \int f(g(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}) + \beta \left( \int g(\boldsymbol{\theta}) dQ(\boldsymbol{\theta}) - 1 \right), \quad (15)$$

$$= -\beta + \int \left( g(\boldsymbol{\theta})(\mathbf{L}_z(\boldsymbol{\theta}) + \beta) + \lambda f(g(\boldsymbol{\theta})) \right) dQ(\boldsymbol{\theta}), \quad (16)$$

where  $\beta$  is a real that acts as a Lagrange multiplier due to the constraint (14b).

Let  $\hat{g} : \mathcal{M} \rightarrow \mathbb{R}$  be a function in  $\mathcal{M}$ . The Gateaux differential of the functional  $L$  in (15) at  $(g, \beta) \in \mathcal{M} \times \mathbb{R}$  in the direction of  $\hat{g}$ , if it exists, is

$$\partial L(g, \beta; \hat{g}) \triangleq \left. \frac{d}{d\gamma} L(g + \gamma \hat{g}, \beta) \right|_{\gamma=0}. \quad (17)$$

Let the function  $r : \mathbb{R} \rightarrow \mathbb{R}$  satisfy for all  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon$  arbitrarily small, that

$$r(\gamma) = L(g + \gamma \hat{g}, \beta) \quad (18)$$

$$\begin{aligned} &= \int \mathbf{L}_z(\boldsymbol{\theta})(g(\boldsymbol{\theta}) + \gamma \hat{g}(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}) + \lambda \int f(g(\boldsymbol{\theta}) + \gamma \hat{g}(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}) \\ &\quad + \beta \left( \int (g(\boldsymbol{\theta}) + \gamma \hat{g}(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}) - 1 \right) \end{aligned} \quad (19)$$

$$\begin{aligned} &= -\beta + \int g(\boldsymbol{\theta})(\mathbf{L}_z(\boldsymbol{\theta}) + \beta) dQ(\boldsymbol{\theta}) + \gamma \int \hat{g}(\boldsymbol{\theta})(\mathbf{L}_z(\boldsymbol{\theta}) + \beta) dQ(\boldsymbol{\theta}) \\ &\quad + \lambda \int f(g(\boldsymbol{\theta}) + \gamma \hat{g}(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}). \end{aligned} \quad (20)$$

Note that the first two terms in (20) are independent of  $\gamma$ ; the third term is linear with  $\gamma$ ; and the fourth term can be written using the function  $\hat{r} : \mathbb{R} \rightarrow \mathbb{R}$  such that for all  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon$  arbitrarily small, it holds that

$$\hat{r}(\gamma) = \lambda \int f(g(\boldsymbol{\theta}) + \gamma \hat{g}(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}). \quad (21)$$

Hence, if the function  $\hat{r}$  in (21) is differentiable at zero, so is the function  $r$  in (18), which implies that the Gateaux differential  $\partial L(g, \beta; \hat{g})$  in (17) exists. The objective is to prove that the function  $\hat{r}$  is differentiable at zero, which boils down to proving that the limit

$$\lim_{\delta \rightarrow 0} \frac{1}{\delta} (\hat{r}(\gamma + \delta) - \hat{r}(\gamma)) \quad (22)$$

exists for all  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon$  arbitrarily small. The proof of the existence of such a limit relies on the fact that from the assumption that  $f$  is strictly convex, it follows that  $f$  is continuous. This implies that  $f$  is also Lipschitz continuous, which implies that for all  $\boldsymbol{\theta} \in \mathcal{M}$  and for all  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon > 0$  arbitrarily small, it holds that

$$|f(g(\boldsymbol{\theta}) + (\gamma + \delta)\hat{g}(\boldsymbol{\theta})) - f(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta}))| \leq c |\hat{g}(\boldsymbol{\theta})| |\delta|, \quad (23)$$

for some constant  $c$  positive and finite, which implies that

$$\left| \frac{f(g(\boldsymbol{\theta}) + (\gamma + \delta)\hat{g}(\boldsymbol{\theta})) - f(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta}))}{\delta} \right| \leq c |\hat{g}(\boldsymbol{\theta})|. \quad (24)$$

Using these arguments, the limit in (22) satisfies for all  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon > 0$  arbitrarily small, that

$$\begin{aligned} & \lim_{\delta \rightarrow 0} \frac{1}{\delta} (\hat{r}(\gamma + \delta) - \hat{r}(\gamma)) \\ &= \lambda \lim_{\delta \rightarrow 0} \frac{1}{\delta} \left( \int f(g(\boldsymbol{\theta}) + (\gamma + \delta)\hat{g}(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}) - \int f(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta})) dQ(\boldsymbol{\theta}) \right) \quad (25) \\ &= \lambda \lim_{\delta \rightarrow 0} \int \frac{f(g(\boldsymbol{\theta}) + (\gamma + \delta)\hat{g}(\boldsymbol{\theta})) - f(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta}))}{\delta} dQ(\boldsymbol{\theta}) \quad (26) \end{aligned}$$

$$= \lambda \int \dot{f}(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta})) \hat{g}(\boldsymbol{\theta}) dQ(\boldsymbol{\theta}) \quad (27)$$

$$< \infty, \quad (28)$$

where both the equality in (27) and the inequality in (28) follow from noticing that the conditions for the dominated convergence theorem hold [35, Theorem 1.6.9], namely:

- For all  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon > 0$ , the inequality in (24) holds;
- The function  $\hat{g}$  in (24) satisfies the inequality in (13); and
- For all  $\boldsymbol{\theta} \in \mathcal{M}$  and for all  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon > 0$  arbitrarily small, it holds that

$$\lim_{\delta \rightarrow 0} \frac{f(g(\boldsymbol{\theta}) + (\gamma + \delta)\hat{g}(\boldsymbol{\theta})) - f(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta}))}{\delta} = \frac{d}{d\gamma} f(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta})) \quad (29)$$

$$= \dot{f}(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta})) \hat{g}(\boldsymbol{\theta}) \quad (30)$$

which follows from the fact that  $f$  is assumed to be differentiable.

From (28), it follows that the function  $\hat{r}$  in (21) is differentiable in the interval  $(-\epsilon, \epsilon)$ , with  $\epsilon > 0$  arbitrarily small. This implies that the function  $r$  in (18) is differentiable within the same interval. The derivative of the real function  $r$  in (18) at  $\gamma \in (-\epsilon, \epsilon)$ , with  $\epsilon > 0$  arbitrarily small, is

$$\frac{d}{d\gamma} r(\gamma) = \int \hat{g}(\boldsymbol{\theta}) (\mathbf{L}_z(\boldsymbol{\theta}) + \beta) dQ(\boldsymbol{\theta}) + \lambda \int \dot{f}(g(\boldsymbol{\theta}) + \gamma\hat{g}(\boldsymbol{\theta})) \hat{g}(\boldsymbol{\theta}) dQ(\boldsymbol{\theta}) \quad (31)$$

$$= \int \hat{g}(\boldsymbol{\theta}) \left( \mathbf{L}_z(\boldsymbol{\theta}) + \beta + \lambda \dot{f}(g(\boldsymbol{\theta})) + \gamma \hat{g}(\boldsymbol{\theta}) \right) dQ(\boldsymbol{\theta}). \quad (32)$$

Hence, the Gateaux differential  $\partial L(g, \beta; \hat{g})$  in (17) exists and satisfies

$$\partial L(g, \beta; \hat{g}) = \left. \frac{d}{d\gamma} r(\gamma) \right|_{\gamma=0} \quad (33)$$

$$= \int \hat{g}(\boldsymbol{\theta}) \left( \mathbf{L}_z(\boldsymbol{\theta}) + \beta + \lambda \dot{f}(g(\boldsymbol{\theta})) \right) dQ(\boldsymbol{\theta}). \quad (34)$$

The relevance of the Gateaux differential in (33) stems from [36, Theorem 1, page 178], which unveils the fact that a necessary condition for the functional  $L$  in (15) to have a stationary point at  $\left( \frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}}{dQ}, \beta \right) \in \mathcal{M} \times [0, +\infty)$  is that for all functions  $\hat{g} \in \mathcal{M}$ ,

$$\partial L \left( \frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}}{dQ}, \beta; \hat{g} \right) = 0. \quad (35)$$

From (34) and (35), it follows that  $\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}}{dQ}$  must satisfy for all functions  $\hat{g}$  in  $\mathcal{M}$  that

$$\int \hat{g}(\boldsymbol{\theta}) \left( \mathbf{L}_z(\boldsymbol{\theta}) + \beta + \lambda \dot{f} \left( \frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta}) \right) \right) dQ(\boldsymbol{\theta}) = 0. \quad (36)$$

This implies that for all  $\boldsymbol{\theta} \in \text{supp } Q$ ,

$$\mathbf{L}_z(\boldsymbol{\theta}) + \beta + \lambda \dot{f} \left( \frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta}) \right) = 0, \quad (37)$$

and thus,

$$\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta}) = \dot{f}^{-1} \left( \frac{-\beta - \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \right), \quad (38)$$

where  $\beta$  is chosen to satisfy (11) and guarantee that for all  $\boldsymbol{\theta} \in \text{supp } Q$ , it holds that  $\dot{f}^{-1} \left( \frac{-\beta - \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \right) \in (0, +\infty)$ . That is,

$$\beta \in \left\{ t \in \mathbb{R} : \forall \boldsymbol{\theta} \in \text{supp } Q, 0 < \dot{f}^{-1} \left( -\frac{t + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \right) < +\infty \right\}, \quad (39)$$

which is an assumption of the theorem.

Finally, note that the objective function in (14) is the sum of two terms. The first one, i.e.,  $\int \mathbf{L}_z(\boldsymbol{\theta}) g(\boldsymbol{\theta}) dQ(\boldsymbol{\theta})$ , is linear with  $g$ . The second, i.e.,  $\int f(g(\boldsymbol{\theta})) dQ(\boldsymbol{\theta})$ , is strictly convex with  $g$  from the assumption that  $f$  is strictly convex. Hence, given that  $\lambda > 0$ , the sum of both terms is strictly convex with  $g$ . This implies the uniqueness of  $P_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}$ . ■

## 4 Examples

This section describes some examples for particular choices of the function  $f$  in Theorem 3.1.

### 4.1 Kullback-Leibler Divergence

Let the function  $f : (0, +\infty) \rightarrow \mathbb{R}$  be such that

$$f(x) = x \log(x), \quad (40)$$

whose derivative satisfies

$$\dot{f}(x) = 1 + \log(x). \quad (41)$$

In this case, the resulting  $f$ -divergence  $D_f(P\|Q)$  is the relative entropy of  $P$  with respect to  $Q$ . From equation (41) and Theorem 3.1, it holds that for all  $\boldsymbol{\theta} \in \text{supp } Q$

$$1 + \log \left( \frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ} \right) = -\frac{\beta + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}, \quad (42)$$

which implies

$$\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ} = \exp \left( -\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \right) \quad (43)$$

$$= \frac{\exp(-\frac{1}{\lambda} \mathbf{L}_z(\boldsymbol{\theta}))}{\int \exp(-\frac{1}{\lambda} \mathbf{L}_z(\nu)) dQ(\nu)}. \quad (44)$$

This result has been independently reported by several authors in [12, 15, 27, 29, 40], and references therein.

### 4.2 Reverse Relative Entropy Divergence

Let the function  $f : (0, +\infty) \rightarrow \mathbb{R}$  be such that

$$f(x) = -\log(x), \quad (45)$$

whose derivative satisfies

$$\dot{f}(x) = -\frac{1}{x}. \quad (46)$$

In this case, the resulting  $f$ -divergence  $D_f(P\|Q)$  is the relative entropy of  $Q$  with respect to  $P$ . From equation (46) and Theorem 3.1, it holds that for all  $\boldsymbol{\theta} \in \text{supp } Q$

$$-\left( \frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ} \right)^{-1} = -\frac{\beta + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}, \quad (47)$$

which implies

$$\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta}) = \frac{\lambda}{\beta + \mathbf{L}_z(\boldsymbol{\theta})}. \quad (48)$$

This result has been reported reported in [14, 41].

### 4.3 Jeffrey's Divergence

Let the function  $f : (0, +\infty) \rightarrow \mathbb{R}$  be such that

$$f(x) = x \log(x) - \log(x), \quad (49)$$

whose derivative satisfies

$$\dot{f}(x) = \log(x) + 1 - \frac{1}{x}. \quad (50)$$

In this case, the resulting  $f$ -divergence  $D_f(P\|Q)$  is Jeffrey's divergence between  $P$  and  $Q$ . From equation (50) and Theorem 3.1, it holds that for all  $\boldsymbol{\theta} \in \text{supp } Q$ ,

$$\log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + 1 - \frac{dQ}{dP_{\Theta|Z=z}^{(Q,\lambda)}}(\boldsymbol{\theta}) = -\frac{\beta + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}, \quad (51)$$

which implies

$$0 = -\frac{dQ}{dP_{\Theta|Z=z}^{(Q,\lambda)}}(\boldsymbol{\theta}) + \log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + 1 + \frac{\beta + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \quad (52)$$

$$= -1 + \frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta}) \left( \log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \right) \quad (53)$$

$$= \exp\left(\log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right)\right) \left( \log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \right) - 1 \quad (54)$$

$$= \left( \exp\left(\log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right)\right) \right) \left( \log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \right) - 1 \frac{\exp\left(\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right)}{\exp\left(\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right)} \quad (55)$$

$$= \exp\left(\log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right) \left( \log\left(\frac{dP_{\Theta|Z=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) \right)$$

$$+\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda} \Big) - \exp\left(\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right). \quad (56)$$

Let  $W_0 : [0, \infty) \rightarrow [0, \infty)$  be the Lambert function, which for a function  $g : \mathbb{R} \rightarrow \mathbb{R}$  such that  $g(x) = x \exp(x)$  satisfies  $W_0(g(x)) = x$ . Hence, from the equality in (56), it holds that

$$\begin{aligned} & \exp\left(\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right) \\ &= \exp\left(\log\left(\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right) \\ & \left(\log\left(\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right), \end{aligned} \quad (57)$$

which in terms of the Lambert function yields

$$W_0\left(\exp\left(\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right)\right) = \log\left(\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})\right) + \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}. \quad (58)$$

Finally, for all  $\boldsymbol{\theta} \in \text{supp } Q$ ,

$$\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta}) = \exp\left(W_0\left(\exp\left(\frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right)\right) - \frac{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}\right). \quad (59)$$

#### 4.4 Hellinger Divergence

Let the function  $f : (0, +\infty) \rightarrow \mathbb{R}$  be such that

$$f(x) = (1 - \sqrt{x})^2, \quad (60)$$

which implies

$$\dot{f}(x) = 1 - \frac{1}{\sqrt{x}}. \quad (61)$$

In this case, the resulting  $f$ -divergence  $D_f(P||Q)$  is the Hellinger's divergence of  $P$  with respect to  $Q$ . From equation (61) and Theorem 3.1, it holds that for all  $\boldsymbol{\theta} \in \text{supp } Q$ ,

$$1 - \frac{1}{\sqrt{\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta})}} = -\frac{\beta + \mathbf{L}_z(\boldsymbol{\theta})}{\lambda}. \quad (62)$$

Hence, from (62) it follows that for all  $\boldsymbol{\theta} \in \text{supp } Q$ ,

$$\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=z}^{(Q,\lambda)}}{dQ}(\boldsymbol{\theta}) = \left(\frac{\lambda}{\beta + \lambda + \mathbf{L}_z(\boldsymbol{\theta})}\right)^2. \quad (63)$$

## 4.5 Jensen-Shannon Divergence

Let the function  $f : (0, +\infty) \rightarrow \mathbb{R}$  be such that

$$f(x) = x \log\left(\frac{2x}{x+1}\right) + \log\left(\frac{2}{x+1}\right), \quad (64)$$

which implies

$$\dot{f}(x) = \log\left(\frac{2x}{x+1}\right). \quad (65)$$

From equation (65) and Theorem 3.1 it holds that for all  $\boldsymbol{\theta} \in \text{supp } Q$ ,

$$\log\left(\frac{2 \frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ}(\boldsymbol{\theta})}{\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ}(\boldsymbol{\theta}) + 1}}\right) = -\frac{\beta + \mathbf{L}_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}, \quad (66)$$

which implies that

$$\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ}(\boldsymbol{\theta}) = \frac{\exp\left(-\frac{\beta + \mathbf{L}_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}\right)}{2 - \exp\left(-\frac{\beta + \mathbf{L}_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}\right)} \quad (67)$$

$$= \frac{1}{2 \exp\left(\frac{\beta + \mathbf{L}_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}\right) - 1}. \quad (68)$$

## 4.6 $\chi^2$ Divergence

Let the function  $f : (0, \infty) \rightarrow \mathbb{R}$  be such that

$$f(x) = (x - 1)^2, \quad (69a)$$

such that

$$\dot{f}(x) = 2(x - 1). \quad (69b)$$

In this case, the resulting  $f$ -divergence  $D_f(P\|Q)$  is the the Pearson-divergence, also known as, the  $\chi^2$ -divergence. From equation (69b) and Theorem 3.1 it holds that for all  $\boldsymbol{\theta} \in \text{supp } Q$

$$2\left(\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ}(\boldsymbol{\theta}) - 1\right) = -\frac{\beta + \mathbf{L}_{\mathbf{z}}(\boldsymbol{\theta})}{\lambda}, \quad (70)$$

which implies

$$\frac{dP_{\boldsymbol{\Theta}|\mathbf{Z}=\mathbf{z}}^{(Q,\lambda)}(\boldsymbol{\theta})}{dQ}(\boldsymbol{\theta}) = \frac{2\lambda - \beta - \mathbf{L}_{\mathbf{z}}(\boldsymbol{\theta})}{2\lambda}. \quad (71)$$

## 5 Conclusions

This work has presented the solution to the ERM- $f$ DR problem under mild conditions on  $f$ , namely, (a) strict convexity; and (b) differentiability. Under these conditions, the optimal measure is shown to be unique and always exists. This result allows obtaining closed-form expressions for ERM problems with regularizations by relative entropy ( $f(x) = x \log(x)$ ), reverse relative entropy ( $f(x) = -\log(x)$ ); Jeffrey's divergence ( $f(x) = x \log(x) - \log(x)$ ); Hellinger divergence ( $f(x) = (1 - \sqrt{x})^2$ ); Jensen-Shannon divergence ( $f(x) = x \log\left(\frac{2x}{x+1}\right) + \log\left(\frac{2}{x+1}\right)$ ); among others. The result is limited in the sense that popular regularizations such as the one by total variation ( $f(x) = |x - 1|$ ) cannot be studied via the main result. This is because in this case,  $f$  is not strictly convex.

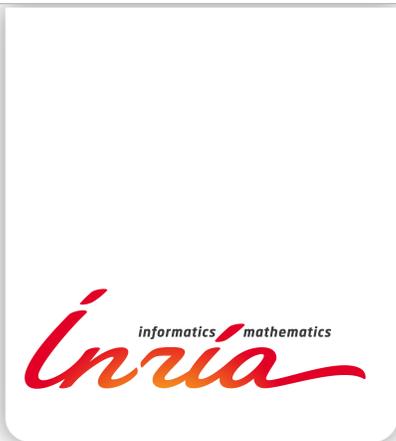
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