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Identifiability of total effects from abstractions of time series causal graphs

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Abstract

We study the problem of identifiability of the total effect of an intervention from observational time series only given an abstraction of the causal graph of the system. Specifically, we consider two types of abstractions: the extended summary causal graph which conflates all lagged causal relations but distinguishes between lagged and instantaneous relations; and the summary causal graph which does not give any indication about the lag between causal relations. We show that the total effect is always identifiable in extended summary causal graphs and we provide necessary and sufficient graphical conditions for identifiability in summary causal graphs. Furthermore, we provide adjustment sets allowing to estimate the total effect whenever it is identifiable.

1 Introduction

Over the last century and across numerous disciplines, experimentation has emerged as a potent methodology for estimating without bias the total effect of an intervention on a specific component of a given system (Neyman et al., 1990). However, experimentation can be costly, unethical or even unfeasible. Both researchers and experts are thus interested in estimating the effect of an intervention directly from observational data. This can be done under

some assumptions when relying on a causal graph (Pearl et al., 2000). Typically, it is tackled through two sequential steps: identifiability and estimation (Pearl, 2019). The identifiability step involves graphically distinguishing cases where a solution is possible and, when it exists, providing an estimand - an expression enabling the estimation of intervention effects from observational data. The subsequent step involves the actual estimation of this estimand from the available data.

The identifiability step received much attention for non-temporal causal graphs (Pearl, 1993, 1995; Spirtes et al., 2000; Pearl et al., 2000; Shpitser and Pearl, 2008; Maathuis and Colombo, 2013; Perkovic, 2020; Anand et al., 2023). Recently, it was also addressed for temporal causal graph. First, Blondel et al. (2016) develop the do-calculus for the full-time causal graphs: each vertex represents a temporal instance of a time series. However, in dynamic systems, experts have difficulties in building time series causal graphs, while they can usually build an abstraction of those graphs where temporal information is omitted. The identifiability problem has only been solved for abstractions of causal graphs either under the assumption that there is no instantaneous relations (Eichler and Didelez, 2007) or under an acyclicity assumption (Assaad et al., 2023).

In this work, we focus on establishing identifiabil-

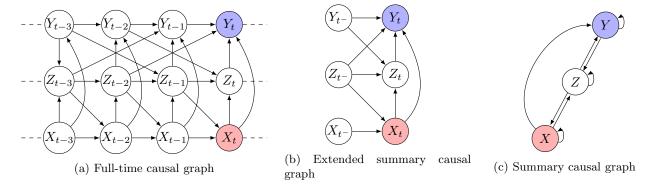


Figure 1: Example of a full-time causal graph (a), its corresponding extended summary causal graph (b) and (c) its corresponding summary causal graph. In (b), t^- represents the past time slice and t the present time slice. In (c), the self loops indicates that a time series causes itself. We illustrate the running example's total effect $P(y_t|do(x_t))$, with $\gamma = 0$, where the blue vertex represents the response and the red vertex the intervention.

ity within abstractions of time series causal graph which may encompass instantaneous causal relationships. Total effects, in time series, take the following form: $P(Y_t = y_t|do(X_{t-\gamma} = x_{t-\gamma}))$, latter denoted $P(y_t|do(x_{t-\gamma}))$ by a slight abuse of notations, where $do(X_{t-\gamma} = x_{t-\gamma})$ represents an intervention (as properly defined in Eichler and Didelez (2007, Assumption 2.3)) on the variable X at time $t - \gamma$, with $\gamma \geq 0$. Here, Y_t corresponds to the response and do() denotes the standard intervention operator (Pearl et al., 2000). Our main contribution lies in demonstrating when, with no hidden confounding (also referred to as causal sufficiency, Spirtes et al. (2000)), the total effect can be identified when working with two specific types of abstraction known as the extended summary causal graph and the summary causal graph.

The remainder of the paper is structured as follows: Section 2 presents the problem setup. In Section 3, we first establish that when using extended summary causal graphs, the total effect is always identifiable, and we provide adjustment sets for its estimation. Section 4 delves into the derivation of both necessary and sufficient conditions for identifiability when employing summary causal graphs, and adjustment sets are outlined for cases where identifiability is achieved. Finally, Section 5 concludes the paper.

2 Problem setup

Consider \mathcal{V} a set of p observational time series and $\mathcal{V}^f = \{\mathcal{V}_{t-\ell} | \ell \in \mathbb{Z}\}$ the set of temporal instances of \mathcal{V} where $\mathcal{V}_{t-\ell}$ correspond to the variables of the time series at time $t-\ell$. We suppose that the time series are generated from a dynamic structural causal model (DSCM, Rubenstein et al. (2018)), an extension of structural causal models (SCM, Pearl et al. (2000)) to time series. The graph that is used to qualitatively represent causal relations described in a DSCM is known as the full-time causal graph (FTCG).

Definition 1 (Full-time causal graph (FTCG), Figure 1a). Let V be a set of p observation time series. The full-time causal graph (FTCG) $\mathcal{G}^f = (V^f, \mathcal{E}^f)$ is defined by set of time series for vertices V^f and the set of edges \mathcal{E}^f defined as follows: $X_{t-\gamma} \to Y_t$ if and only if X causes Y at time t with a time lag of $\gamma > 0$ if X = Y and with a time lag of $\gamma \geq 0$ for $X \neq Y$.

The FTCG is assumed to be acyclic and to satisfy the Markov condition. However, in practice, it is usually impossible to work with FTCGs and people have resorted to simpler causal graphs, exploiting the fact that causal relations between time series hold throughout time, as formalized in the following assumption.

Assumption 1 (Consistency throughout time). An FTCG \mathcal{G}^f is said to be consistent throughout time if all the causal relationships remain constant in direction throughout time.

This assumption allows to focus on a finite number of past slices, which is given by the maximum lag. We fix it to $\gamma_{\rm max}$ in the following.

Experts are used to working with abstractions of causal graphs, that summarize the information into a graph that is interpretable, often with the omission of precise temporal information details. An extended summary causal graph (Assaad et al., 2022a) serves as a representation for causal relationships among time series. It distinguishes between past time slices, denoted as $\mathcal{V}^e_{t^-}$, and present time slices, denoted as $\mathcal{V}^e_{t^-}$, thus enabling the differentiation between lagged and instantaneous causal relations.

Definition 2 (Extended summary causal graph (ECG), Figure 1b). Let $\mathcal{G}^f = (\mathcal{V}^f, \mathcal{E}^f)$ be an FTCG satisfying Assumption 1 with maximal temporal lag γ_{\max} . The underlying extended summary causal graph (ECG) $\mathcal{G}^e = (\mathcal{V}^e, \mathcal{E}^e)$ is given by $\mathcal{V}^e = (\mathcal{V}^e_{t-}, \mathcal{V}^e_{t})$ and \mathcal{E}^e defined as follows:

- for all $X_t, Y_t \in \mathcal{V}_t^e$, $X_t \to Y_t \in \mathcal{E}^e$ if and only if $X_t \to Y_t \in \mathcal{E}^f$;
- for all $X, Y \in \mathcal{V}_{t^-}^e$, $X_{t^-} \to Y_t \in \mathcal{E}^e$ if and only if there exists at least one temporal lag $0 < \gamma \le \gamma_{\max}$ such that $X_{t-\gamma} \to Y_t \in \mathcal{E}^f$.

In that case, we say that \mathcal{G}^f is compatible with \mathcal{G}^e .

At a higher level of abstraction, a summary causal graph (Peters et al., 2013; Meng et al., 2020; Assaad et al., 2022b,c) represents causal relationships among time series, regardless of the time delay between the cause and its effect.

Definition 3 (Summary causal graph (SCG), Figure 1c). Let $\mathcal{G}^f = (\mathcal{V}^f, \mathcal{E}^f)$ be an FTCG satisfying Assumption 1 with maximal temporal lag γ_{max} . The underlying summary causal graph (SCG) $\mathcal{G}^s = (\mathcal{V}^s, \mathcal{E}^s)$ is such that

• V^s corresponds to the set of time series,

• $X \to Y \in \mathcal{E}^s$ if and only if there exists at least one temporal lag $0 \le \gamma \le \gamma_{\max}$ such that $X_{t-\gamma} \to Y_t \in \mathcal{E}^f$.

In that case, we say that \mathcal{G}^f is compatible with \mathcal{G}^s .

The abstractions of SCGs and ECGs entail that if there is a single ECG or SCG corresponding to a given FTCG, there are often multiple FTCGs compatible with a given ECG or SCG. Additionally, it is worth noting that an SCG is considered an abstraction of an ECG. However, in cases where there is no instantaneous relation, ECGs and SCGs convey the same information.

Since the FTCG is assumed to be acyclic, the ECG is inherently acyclic. In contrast, SCGs may include directed cycles and even self-loops. We use the notation $X \rightleftharpoons Y$ to indicate situations where there are time lags where X causes Y and other lags where Y causes X. We denote $Cycles(X, \mathcal{G}^s)$ as the set of all directed cycles containing X in \mathcal{G}^s , and $Cycles^>(X,\mathcal{G}^s)$ with at least 2 vertices (i.e., excluding self-loops). By a slight abuse of notation, we denote $\mathcal{G}^s \setminus \{Y\}$ as the subgraph of \mathcal{G}^s when removing the vertex Y and corresponding edges.

We focus in this paper on the total effect (Pearl et al., 2000) of $X_{t-\gamma}$ on Y_t given by $P(y_t|do(x_{t-\gamma}))$, the distribution of Y_t while $X_{t-\gamma}$ is held constant at $x_{t-\gamma}$ and all other variables are permitted to run their natural course according to the DSCM. The goal of this paper is to discuss the identifiability of the total effect given an ECG or an SCG, without the knowledge of the true FTCG. In general, a total effect $P(y_t \mid do(x_{t-\gamma}))$ is said to be identifiable from a graph if it can be uniquely computed from the observed distribution, without any further assumption on the distribution (Pearl, 1995; Perkovic, 2020). More specifically, in our context, $P(y_t \mid do(x_{t-\gamma}))$ is identifiable if it can be rewritten with a do-free formula that is valid for all FTCGs compatible with the given ECG or SCG. This consists on finding for all FTCGs a set of variables known as adjustment set (Shpitser et al., 2010) which permits to block all non-causal paths without blocking any causal path between $X_{t-\gamma}$ and Y_t . This leads to the following graphical definition of identifiability.

Definition 4 (Identifiability of total effects in ECGs and SCGs). With no hidden confounding, given an ECG or an SCG, the total effect $P(y_t|do(x_{t-\gamma}))$ is identifiable if and only if there exists a set that:

- (i) blocks all back-door paths between $X_{t-\gamma}$ and Y_t in all compatible FTCGs,
- (ii) does not block any directed path between $X_{t-\gamma}$ and Y_t in all compatible FTCGs,
- (iii) does not contain any descendant of Y_t in all compatible FTCGs.

This set is called an adjustment set for $P(y_t|do(x_{t-\gamma}))$ in the ECG or the SCG.

Note that, without hidden confounding, the total effect $P(y_t \mid do(x_{t-\gamma}))$ in any FTCG \mathcal{G}^f is identifiable (Blondel et al., 2016; Shpitser and Pearl, 2008). However, since an ECG or an SCG usually represents different possible FTCGs, it is not a priori certain that a given total effect $P(y_t | do(x_{t-\gamma}))$ will yield equivalent do-free formulas across the various FTCGs.

3 Identifiability in ECGs

With no hidden confounding, the total effect is always identifiable in ECGs, as established by the following theorem.

Theorem 1 (Identifiability in ECG). Consider an ECG \mathcal{G}^e . The total effect $P(y_t|do(x_{t-\gamma}))$, with $\gamma \geq 0$ is identifiable in \mathcal{G}^e .

Proof. If $X \notin Anc(Y, \mathcal{G}^e)$, then in every FTCG \mathcal{G}^f compatible with \mathcal{G}^e , $X_{t-\gamma} \notin Anc(Y_t, \mathcal{G}^f)$. Thus, $P(y_t \mid do(x_{t-\gamma}))$ is always identifiable in \mathcal{G}^e , and $P(y_t \mid do(x_{t-\gamma})) = P(y_t)$.

Assume now that $X \notin Anc(Y, \mathcal{G}^e)$. Let γ_{\max} be the maximal lag, and \mathcal{G}^f be an FTCG compatible with \mathcal{G}^e . We prove that

$$\mathcal{B}_{\gamma} = \{ (Z_{t-\gamma-\ell})_{1 \le \ell \le \gamma_{\max}} | Z_{t^-} \in Par(X_t, \mathcal{G}^e) \}$$
$$\cup \{ Z_{t-\gamma} | Z_t \in Par(X_t, \mathcal{G}^e) \}$$

is an adjustment set for $P(y_t|do(x_{t-\gamma}))$ in \mathcal{G}^f .

First, we have to prove that $Par(X_{t-\gamma}, \mathcal{G}^f) \subseteq \mathcal{B}_{\gamma}$. Let $Z_{t-\gamma-\ell} \in Par(X_{t-\gamma}, \mathcal{G}^f)$. If $\ell = 0$, then Z_t causes X_t in \mathcal{G}^e by consistency throughout time, which means that $Z_{t-\gamma} \in \mathcal{B}_{\gamma}$. If $\ell > 0$, then Z_{t-} causes X_t in \mathcal{G}^e , that is $Z_{t-\ell-\gamma} \in \mathcal{B}_{\gamma}$. This shows that the set \mathcal{B}_{γ} blocks all back-door paths relatively to $P(y_t|do(x_{t-\gamma}))$.

Then, we have to prove \mathcal{B}_{γ} does not contain any descendant of $X_{t-\gamma}$ in \mathcal{G}^f . If this is true, there exists $W_{t-\gamma} \in \mathcal{B}_{\gamma} \cap Desc(X_{t-\gamma}, \mathcal{G}^f)$, at time slice $t-\gamma$ because it is a parent and a descendant of $X_{t-\gamma}$. By consistency throughout time, $W_t \in Desc(X_t, G^f)$. However, by definition of \mathcal{B}_{γ} , $X_t \in Desc(W_t, G^f)$, which contradicts the acyclicity assumption of \mathcal{G}^e . It means that neither it blocks any directed path between $X_{t-\gamma}$ and Y_t , nor it contains any descendant of Y_t .

As every total effect is identifiable in ECG, it is relevant to determine an adjustment set. The set \mathcal{B}_{γ} , as defined in the proof of Theorem 1, can serve as an adjustment set for $P(y_t|do(x_{t-\gamma}))$, though this set is very large. The following proposition gives a range of adjustment sets, which are potentially smaller. This proposition is based on the back-door criterion of the densest compatible FTCG, defined as follows: for an ECG \mathcal{G}^e and a maximal lag γ_{max} , the densest compatible FTCG is the FTCG compatible with \mathcal{G}^e with all potential edges.

Proposition 1. Consider an ECG \mathcal{G}^e and a maximal lag γ_{\max} and let $\gamma \geq 0$. Any adjustment set \mathcal{B}'_{γ} for the total effect $P(y_t|do(x_{t-\gamma}))$ given by the back-door criterion on the densest compatible FTCG is a valid adjustment set for the total effect in \mathcal{G}^e and

$$P(y_t|do(x_{t-\gamma})) = \sum_{\mathbf{b} \in \mathcal{B}'_{\gamma}} P(\mathbf{b}) P(y_t|x_{t-\gamma}, \mathbf{b}),$$

where the sum covers all values **b** in the support of \mathcal{B}'_{γ} .

Proof. Let \mathcal{G}_d^f be the densest FTCG compatible with \mathcal{G}^s , and \mathcal{B}_{γ}^f be an adjustment set in \mathcal{G}_d^f . Let \mathcal{G}^f be another FTCG compatible with \mathcal{G}^e . By definition of \mathcal{G}_d^f , any back-door path in \mathcal{G}^f is also in \mathcal{G}_d^f (the last graph contains all possible edges). Then, \mathcal{B}_{γ}^f blocks

all back-door paths in \mathcal{G}^f . Moreover, since no vertex in \mathcal{B}'_{γ} is a descendant of $X_{t-\gamma}$ in \mathcal{G}^f_d , the same holds for \mathcal{G}^f . Thus, \mathcal{B}'_{γ} is also an adjustment set in \mathcal{G}^f . \square

Note that \mathcal{B}_{γ} is in this set and that this adjustment set \mathcal{B}'_{γ} is distinct from the back-door criterion for the total effect $P(y_t|do(x_{t-\gamma}))$ as applied to the true FTCG, which we do not know.

4 Identifiability in SCGs

SCGs are the most abstract version of causal graphs considered in this paper. Due to the potential cycles, we use the notion of σ -blocked path introduced in Forré and Mooij (2017) but we restrict it to the case of no hidden confounding.

Definition 5 (σ -blocked path with no hidden confounding). Consider an SCG $\mathcal{G}^s = (\mathcal{V}^s, \mathcal{E}^s)$. A path $\pi = \langle X, \dots, Y \rangle$ in \mathcal{G}^s is σ -blocked by a set $\mathcal{Z} \subseteq \mathcal{V}^s$ if:

- at least one of the endnodes X, Y is in Z; or
- there exists a vertex $Z \in \pi$, such that $Z \notin Anc(\mathcal{Z}, \mathcal{G}^s)$ with two adjacent edges in π that form a collider at Z, (i.e., $\to Z \leftarrow$); or
- there exists a vertex $Z \in \pi \cap \mathcal{Z}$ with two adjacent edges in π that forms a non-collider at Z, and such that there exists a directed edge pointing from Z to a vertex not in the strongly connected component $Sc(Z, \mathcal{G}^s) = Anc(Z, \mathcal{G}^s) \cap Desc(Z, \mathcal{G}^s)$.

If none of the above holds then the path $\langle X, \dots, Y \rangle$ is σ -active by \mathcal{Z} .

Note that a back-door path can start with the edge \rightleftarrows , and particularly $X \rightleftarrows Y$ is a σ -active back-door path in \mathcal{G}^s .

Given the above definition, we can now present the main result of this section, which states necessary and sufficient conditions for the identifiability in SCG.

Theorem 2 (Identifiability in SCG). Consider an $SCG \mathcal{G}^s = (\mathcal{V}^s, \mathcal{E}^s)$. The total effect $P(y_t|do(x_{t-\gamma}))$, with $\gamma \geq 0$, is not identifiable if and only if $X \in Anc(Y, \mathcal{G}^s)$ and one of the following holds:

- $\gamma \neq 0$ and $Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}) \neq \emptyset$, or
- there exists a σ -active back-door path

$$\pi^s = \langle V^1 = X, \cdots, V^n = Y \rangle$$

from X to Y in \mathcal{G}^s such that $\langle V^2, \cdots, V^{n-1} \rangle \subseteq Desc(X, \mathcal{G}^s)$ and one of the following holds:

- -n > 2, ie $\langle V^2, \cdots, V^{n-1} \rangle \neq \emptyset$, or
- $-\gamma \neq 1$, or
- $\gamma = 1, n = 2 \text{ and } Cycles(Y, \mathcal{G}^s \setminus \{X\}) \neq \emptyset.$

Proof. To prove this theorem, we present necessary conditions in Lemmas 1-4 and sufficient conditions for the identifiability in Lemmas 5-7.

We first introduce the notion of ambiguous vertices that will be useful for the proofs of most of the lemmas.

Definition 6 (Ambiguous vertices). Consider an $SCG \mathcal{G}^s$ and the total effect $P(y_t \mid do(x_{t-\gamma}))$, for $\gamma \geq 0$. A vertex $V_{t'}$ is ambiguous for $P(y_t \mid do(x_{t-\gamma}))$ if there exist two $FTCGs \mathcal{G}_1^f$ and \mathcal{G}_2^f compatible with \mathcal{G}^s such that:

- $V_{t'}$ belongs to a back-door path of $(X_{t-\gamma}, Y_t)$ in \mathcal{G}_1^f ; and
- $V_{t'}$ belongs to a directed path from $X_{t-\gamma}$ to Y_t in \mathcal{G}_2^f or $V_{t'} \in Desc(Y_t, \mathcal{G}_2^f)$.

Ambiguous vertices are the one that prevent identifiability, as shown in the next property.

Property 1. Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-\gamma}))$ for $\gamma \geq 0$. If there exists a backdoor path π^f in an FTCG compatible with \mathcal{G}^s between $X_{t-\gamma}$ and Y_t in which all vertices are ambiguous then $P(y_t \mid do(x_{t-\gamma}))$ is not identifiable.

Proof. Let's assume that the total effect is identifiable. Then, there exists a set of vertices W which blocks all back-door paths between $X_{t-\gamma}$ and Y_t of all FTCGs compatible with \mathcal{G}^s and which does not block any directed paths between $X_{t-\gamma}$ and Y_t nor contains any descendant of Y_t . Let's consider an FTCG compatible with \mathcal{G}^s and a back-door path π^f between

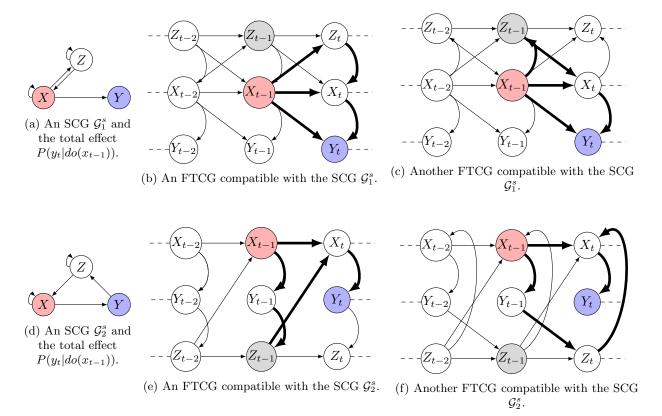


Figure 2: An example of an SCG \mathcal{G}_1^s in (a) (resp. \mathcal{G}_2^s in (d)) with two of its compatible FTCGs in (b) and (c) (resp. in (e) and (f)). Each pair of red and blue vertices in the FTCGs represents one total effect we are interested in. Gray vertices are ambiguous, and bold edges correspond to direct path from $X_{t-\gamma}$ to Y_t . In both examples, the total effect is not identifiable given the SCG as stated in Lemma 1 (resp. Lemma 2).

 $X_{t-\gamma}$ and Y_t in this FTCG. All vertices which are in \mathcal{W} and π^f block π^f . As they are in \mathcal{W} , they cannot belong, in any FTCG compatible with \mathcal{G}^s , to a directed path from $X_{t-\gamma}$ and Y_t , as otherwise they would block this path, and cannot be a descendant of Y_t . They are thus unambiguous according to Definition 6.

This enables us to prove that the conditions stated in Theorem 2 are necessary. We detail in Lemmas 3 and 4 all the cases, and depict them in Figures 2-3. We present an SCG and two FTCGs compatible with the SCG that do not share any adjustment set. Ambiguous vertices are highlighted in gray. The proofs

of the lemmas follow a similar approach: we provide two FTCGs that are compatible with the SCG, along with associated paths in which all vertices are ambiguous. Thus, by Property 1, the total effect is not identifiable.

Lemma 1. Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-\gamma}))$ for $\gamma > 0$. If $X \in Anc(Y, \mathcal{G}^s)$ and $Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}) \neq \emptyset$ then $P(y_t \mid do(x_{t-\gamma}))$ is not identifiable.

This lemma is illustrated in Figure 2 (top row) with the SCG \mathcal{G}_1^s and the total effect $P(y_t|do(x_{t-1}))$.

Proof. Let $\pi^s = X \to W^2 \to \ldots \to W^{m-1} \to Y$ be a

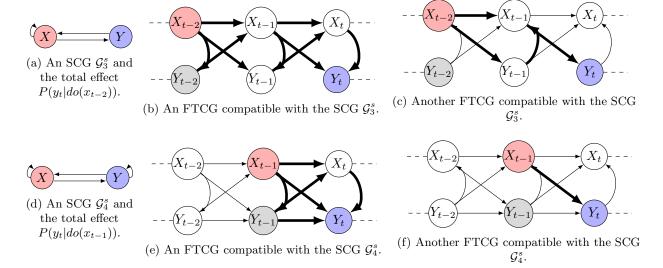


Figure 3: An example of an SCG \mathcal{G}_3^s in (a) (resp. \mathcal{G}_4^s in (d)) with two of its compatible FTCGs in (b) and (c) (resp. in (e) and (f)). Each pair of red and blue vertices in the FTCGs represents one total effect we are interested in. Gray vertices are ambiguous, and bold edges correspond to direct path from $X_{t-\gamma}$ to Y_t . In both examples, the total effect is not identifiable given the SCG as stated in Lemma 3 (resp. Lemma 4).

directed path in \mathcal{G}^s and $C = \langle X, V^2, \dots, V^{n-1}, X \rangle \in Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}).$

There exists an FTCG \mathcal{G}_1^f compatible with \mathcal{G}^s such that $\pi_1^f = X_{t-\gamma} \to \dots \to V_{t-\gamma}^{n-1} \to X_{t-\gamma+1} \to W_{t-\gamma+1}^2 \dots \to Y_t$. There also exists an FTCG \mathcal{G}_2^f compatible with \mathcal{G}^s with the path $\pi_2^f = X_{t-\gamma} \leftarrow V_{t-\gamma}^{n-1} \leftarrow \dots X_{t-\gamma+1} \to \dots \to Y_t$. Thus, by Property 1, $P(y_t \mid do(x_{t-\gamma}))$ is not identifiable.

Lemma 2. Consider an SCG \mathcal{G}^s , and the total effect $P(y_t \mid do(x_{t-\gamma}))$, for $\gamma \geq 0$. If $X \in Anc(Y, \mathcal{G}^s)$ and there exists a σ -active back-door path

$$\pi^s = \langle V^1 = X, \cdots, V^n = Y \rangle$$

from X to Y in \mathcal{G}^s such that $\langle V^2, \cdots, V^{n-1} \rangle \subseteq Desc(X, \mathcal{G}^s)$ and n > 2 then $P(y_t \mid do(x_{t-\gamma}))$ is not identifiable.

This lemma is illustrated in Figure 2 (bottom row) with the SCG \mathcal{G}_2^s and the total effect $P(y_t|do(x_{t-1}))$.

Proof. Let $\pi^s = \langle V^1 = X, \cdots, V^n = Y \rangle$ be such a σ -active backdoor path.

Then there exists an FTCG \mathcal{G}_1^f compatible with \mathcal{G}^s such that $\pi_1^f = X_{t-\gamma} \leftarrow \ldots \leftarrow V_{t-\gamma}^i \rightarrow \ldots \rightarrow Y_t$ is a backdoor path. As $X \in Anc(Y, \mathcal{G}^s)$ and all $V^j \subseteq Desc(X, \mathcal{G}^s)$, $2 \leq j \leq n-1$, all those variables also belong to a direct path from $X_{t-\gamma}$ to Y_t . Thus, by Property 1, $P(y_t \mid do(x_{t-\gamma}))$ is not identifiable. \square

Lemma 3. Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-\gamma}))$, with $\gamma \neq 1$. If $X \leftrightarrows Y \in \mathcal{G}^s$ then $P(y_t \mid do(x_{t-\gamma}))$ is not identifiable.

This lemma is illustrated in Figure 3 (top row) with the SCG \mathcal{G}_3^s and the total effect $P(y_t|do(x_{t-2}))$.

Proof. If $\gamma = 0$, there exists and FTCG \mathcal{G}_1^f compatible with \mathcal{G}^s such that $X_t \in Par(Y_t, \mathcal{G}_1^f)$ and another FTCG \mathcal{G}_2^f compatible with \mathcal{G}^s such that $Y_t \in Par(X_t, \mathcal{G}_2^f)$, thus the total effect is not identifiable.

Let $\gamma > 1$. There exists an FTCG \mathcal{G}_1^f compatible with \mathcal{G}^s such that $X_{t-\gamma} \leftarrow Y_{t-\gamma} \to X_{t-\gamma+1} \to \dots \to Y_t$ and $X_{t-\gamma} \to Y_t$. There also exists an FTCG \mathcal{G}_2^f compatible with \mathcal{G}^s such that $X_{t-\gamma} \to Y_t$.

 $Y_{t-\gamma} \to X_{t-\gamma+1} \to \ldots \to Y_t$. Then, by Property 1, $P(y_t \mid do(x_{t-\gamma}))$ is not identifiable.

Lemma 4. Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-1})) \ (\gamma = 1).$ If $X \subseteq Y \in \mathcal{G}^s$ and $Cycles(Y, \mathcal{G}^s \setminus \{X\}) \neq \emptyset$ then $P(y_t \mid do(x_{t-1}))$ is not identifiable.

This lemma is illustrated in Figure 3 (bottom row) with the SCG \mathcal{G}_4^s and the total effect $P(y_t|do(x_{t-1}))$.

and a directed path $X_{t-1} \to Y_t$. There also exists an FTCG \mathcal{G}_2^f compatible with \mathcal{G}^s with a directed $X_{t-1} \to Y_{t-1} \to \dots \to V_{t-1}^{n-1} \to Y_t$. Then, by Property 1, $P(y_t|do(x_{t-1}))$ is not identifiable.

Now, we turn to Lemmas 5-7, that show that conditions in Theorem 2 are sufficient. To do so, we consider the following set:

$$\mathcal{A}_{\gamma} = \{ (Z_{t-\gamma-\ell})_{1 \leq \ell \leq \gamma_{\max}} | Z \in Desc(X; \mathcal{G}^s) \}$$

$$\cup \{ (Z_{t-\gamma-\ell})_{0 \leq \ell \leq \gamma_{\max}} | Z \in \mathcal{V}^s \backslash Desc(X, \mathcal{G}^s) \}$$

and we prove that it is an adjustment set in each case. As one can note, it contains all possible parents of $X_{t-\gamma}$ in all FTCGs \mathcal{G}^f compatible with \mathcal{G}^s . Thus, \mathcal{A}_{γ} blocks any back-door path π between $X_{t-\gamma}$ and Y_t in any FTCG \mathcal{G}^f compatible with \mathcal{G}^s through the parent of $X_{t-\gamma}$ on that path (which is the first node on the path from $X_{t-\gamma}$ to Y_t), which cannot be a

In addition to ambiguous vertices, one can also define ambiguous paths, as follows.

Definition 7 (Ambiguous paths). Consider an SCG \mathcal{G}^s and an FTCG \mathcal{G}^f compatible with \mathcal{G}^s . A path $\pi^f \in$ \mathcal{G}^f between $X_{t-\gamma}$ and Y_t , for $\gamma \geq 0$, is an ambiguous path if it does not contain any vertex at time $t-\gamma-\ell$ for $\ell > 1$. We note Π_{γ}^f the set of all ambiguous paths in \mathcal{G}^f .

When π is not in Π_{γ}^f , then the parent of $X_{t-\gamma}$ is in the past of $X_{t-\gamma}$ and thus cannot be ambiguous. One thus has the following property:

Property 2. Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-\gamma}))$, for $\gamma \geq 0$. Suppose π^f is a back-door path between $X_{t-\gamma}$ and Y_t in a given FTCG \mathcal{G}^f compatible with \mathcal{G}^s . If $\pi^f \notin \Pi^f_{\infty}$, then π^f is blocked by a subset of A_{γ} containing at least one non-ambiguous vertex.

We now introduce the notion of compatible path.

Definition 8 (Compatible path). Consider an SCG \mathcal{G}^s , an FTCG \mathcal{G}^f compatible with \mathcal{G}^s , and the to-Proof. Let $\langle Y, V^2, \dots, V^{n-1}, Y \rangle \in Cycles(Y, \mathcal{G}^s \setminus \{X\})$ tal effect $P(y_t \mid do(x_{t-\gamma}))$, for $\gamma \geq 0$. We say There exists an FTCG \mathcal{G}_1^f compatible with \mathcal{G}^s with that a path $\pi^f = \langle X_{t-\gamma}, W_{t^2}^2, \dots, W_{t^{m-1}}^{m-1}, Y_t \rangle$ in \mathcal{G}^f is a back-door path $X_{t-1} \leftarrow Y_{t-1} \rightarrow \dots \rightarrow V_{t-1}^{n-1} \rightarrow Y_t$ compatible with a path $\pi^s = \langle X, V^2, \dots, V^{n-1}, Y \rangle$ in \mathcal{G}^s if for all $(W^j_{t^j})_{2 \leq j \leq m-1}$: either $W^j \in \langle V^2, \cdots, V^{n-1} \rangle$ or $\exists V \in \langle V^2, \cdots, V^{n-1} \rangle$ such that $W^j \in Cycles(V, \mathcal{G}^s) \backslash Cycles(X, \mathcal{G}^s)$.

> The following property characterizes any back-door path in an FTCG compatible with the SCG through graphical conditions on the SCG.

> **Property 3.** Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-\gamma}))$ for $\gamma \geq 0$. Then $(i) \Rightarrow (ii)$, where:

- (i) $\gamma = 0$ or $Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}) = \emptyset$,
- (ii) in an FTCG \mathcal{G}^f compatible with \mathcal{G}^s , there exists no back-door path $\pi^f \in \Pi^f_{\gamma}$ that is not compatible with any back-door path in \mathcal{G}^s .

Proof. Assume first $Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}) = \emptyset$. Suppose $\exists \pi^f = X_{t-\gamma} \leftarrow W_{t-\gamma} \cdots \rightarrow Y_t \in \Pi^f_{\gamma}$ which is a back-door path between $X_{t-\gamma}$ and Y_t that is not compatible with any back-door path $\pi^s = \langle V^1 =$ $X, V^2, \cdots, V^{n-1}, V^n = Y$ in \mathcal{G}^s .

If n = 2, then the path compatible with the cycle $\langle X, X \rangle$ is of the form $X_{t-\gamma} \to X_{t-\gamma+i} \to \cdots \to$ $X_{t-\gamma+j} \to Y_t$: it means that π^f cannot be a backdoor path.

If n > 2, $W_{t-\gamma}$ is such that $W \notin \{V^2, \dots, V^{n-1}\}$ and $\nexists V \in \{V^2, \cdots, V^{n-1}\}$ such that $W \in$ $Cycles(V, \mathcal{G}^s)$. If the path between $W_{t-\gamma}$ and Y_t does not pass by $X_{t-\gamma+\ell}$ with $\ell>0$, then there exists a back- door path between X and Y passing by W in \mathcal{G}^s as π^f lies in an FTCG compatible with \mathcal{G}^s , which contradicts our assumption. So the path necessarily

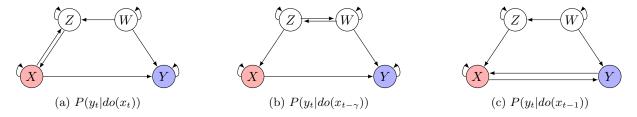


Figure 4: Three SCGs and a total effect which is identifiable. Each pair of red and blue vertices in the FTCGs represents the total effect we are interested in, and we precise the total effect and the lag considered in the caption. This illustrates Lemma 6 (Figure a-b) and Lemma 7 (Figure c).

passes by $X_{t-\gamma+\ell}$. Thus there is a cycle C_x on X such that $size(C_x) > 2$, which leads again to a contradiction. Thus, there does not exist a back-door path $\pi^f \in \Pi^f_\gamma$ between $X_{t-\gamma}$ and Y_t that is not compatible with any back-door path in \mathcal{G}^s .

The case $\gamma = 0$ is treated in the same way, with the fact that the path considered cannot go back to X_t as this would create a cycle in the FTCG.

In the following, we prove the lemmas that correspond to the identifiability cases.

Lemma 5. Consider an SCG \mathcal{G}^s , $\gamma \geq 0$ fixed and the total effect $P(y_t \mid do(x_{t-\gamma}))$. If $X \notin Anc(Y, \mathcal{G}^s)$ then $P(y_t \mid do(x_{t-\gamma}))$ is identifiable.

Proof. If $X \notin Anc(Y, \mathcal{G}^s)$, then in every FTCG \mathcal{G}^f compatible with \mathcal{G}^s , $X_{t-\gamma} \notin Anc(Y_t, \mathcal{G}^f)$. Thus, $P(y_t \mid do(x_{t-\gamma}))$ is always identifiable in \mathcal{G}^s , and $P(y_t \mid do(x_{t-\gamma})) = P(y_t)$.

Lemma 6. Consider an SCG \mathcal{G}^s , $\gamma \geq 0$ fixed and the total effect $P(y_t \mid do(x_{t-\gamma}))$. If $X \in Anc(Y, \mathcal{G}^s)$ and

- (i) either $\gamma = 0$ or $Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}) = \emptyset$ and
- (ii) $\nexists \sigma$ -active back-door path $\pi^s = \langle V^1 = X, \cdots, V^n = Y \rangle$ from X to Y in \mathcal{G}^s such that $\langle V^2, \cdots, V^{n-1} \rangle \subseteq Desc(X, \mathcal{G}^s),$

then $P(y_t \mid do(x_{t-\gamma}))$ is identifiable.

This lemma is illustrated in Figure 4a - 4b.

Proof. We will prove that A_{γ} is an adjustment set for $P(y_t \mid do(x_{t-\gamma}))$ in any FTCG compatible with

 \mathcal{G}^s under conditions (i) and (ii). Let \mathcal{G}^f be an FTCG compatible with \mathcal{G}^s , and Π^f_{γ} the set of ambiguous paths. By Property 2, any back-door path $\pi^f \notin \Pi^f_{\gamma}$ can be blocked by \mathcal{A}_{γ} . Furthermore, by definition, elements of \mathcal{A}_{γ} cannot block any directed path between $X_{t-\gamma}$ and Y_t and cannot be descendants of Y_t as otherwise they would be descendant of $X_{t-\gamma}$.

We now turn our attention to paths in Π_{∞}^f . Let $\pi^f \in \Pi^f_{\gamma}$ be a back-door path between $X_{t-\gamma}$ and Y_t . Since $\gamma = 0$ or $Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}) = \emptyset$ then by Property 3, all back-door paths in Π_{α}^f are compatible with back-door paths in \mathcal{G}^s . Let $\pi^s =$ $\langle V^1 = X, \cdots, V^n = Y \rangle$ be a σ -active back-door path in \mathcal{G}^s compatible with π^f . By (ii), there exists $m \geq 1$ vertices such that $\{V^{i_1}, \dots, V^{i_m}\} \subseteq$ $\langle V^2, \cdots, V^{n-1} \rangle$ and $\{V^{i_1}, \cdots, V^{i_m}\} \not\subset Desc(X, \mathcal{G}^s)$. Then, $\forall V_{t-\gamma}$ such that $V \in \{V^{i_1}, \dots, V^{i_m}\}, V_{t-\gamma} \notin$ $Desc(X_{t-\gamma},\mathcal{G}^f)$ and since $X \in Anc(Y,\mathcal{G}^s)$ then it must be the case that $V \not\in Desc(Y, \mathcal{G}^s)$ and by consequence $V_{t-\gamma} \notin Desc(Y_t, \mathcal{G}^f)$. Thus, $V_{t-\gamma}$ cannot be an ambiguous vertex. Its parent in π^f furthermore blocks π^f , is not ambiguous (as otherwise $V_{t-\gamma}$ would be ambiguous) and is a member of A_{γ} by definition of \mathcal{A}_{γ} . Thus \mathcal{A}_{γ} blocks all back-door paths between $X_{t-\gamma}$ and Y_t in any \mathcal{G}^f compatible with \mathcal{G}^s . Furthermore, no node in A_{γ} can block a directed path between $X_{t-\gamma}$ and Y_t or is a descendant of Y_t as nodes in \mathcal{A}_{γ} are either defined before $t-\gamma$ or are not descendant of $X_{t-\gamma}$, and thus of Y_t . This concludes the

Lemma 7. Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-1}))$ $(\gamma = 1)$. If the only σ -active backdoor path from X to Y in \mathcal{G}^s is $X \subseteq Y \in \mathcal{G}^s$ and

- (i) $Cycles(Y, \mathcal{G}^s \setminus \{X\}) = \emptyset$ and
- (ii) $Cycles^{>}(X, \mathcal{G}^s \setminus \{Y\}) = \emptyset$,

then $P(y_t \mid do(x_{t-1}))$ is identifiable.

This lemma is illustrated in Figure 4c.

Proof. We will prove that \mathcal{A}_1 is an adjustment set for $P(y_t \mid do(x_{t-1}))$ in any FTCG compatible with \mathcal{G}^s under conditions (i) and (ii). Let \mathcal{G}^f be an FTCG compatible with \mathcal{G}^s , and Π_1^f the set of ambiguous paths.

Since (ii) then by Property 3, all back-door paths in Π_1^f are compatible with back-door paths in \mathcal{G}^s . In addition, by Property 2, any path $\pi^f \notin \Pi_1^f$ can be blocked by \mathcal{A}_1 . Therefore, in the following, we focus on paths in Π_1^f compatible with back-door paths in \mathcal{G}^s .

Consider the σ -active back-door path $\pi^s = \langle X, Y \rangle$. As there cannot be a loop on Y by (i), the only path $\pi^f \in \Pi_1^f$ from X_{t-1} to Y_t compatible with π^s that pass by Y_{t-1} is $\pi_f = \langle X_{t-1}, Y_{t-1}, X_t, Y_t \rangle$. Then, under consistency throughout time, acyclicity and temporal priority, the only choices are $X_{t-1} \to Y_{t-1} \to X_t \to Y_t$ and $X_{t-1} \leftarrow Y_{t-1} \to X_t \leftarrow Y_t$. The first is a directed path, the second a back-door path already blocked due to the collider $Y_{t-1} \to X_t \leftarrow Y_t$. Thus, all potential back-door paths between X_{t-1} and Y_t in any \mathcal{G}^f compatible with \mathcal{G}^s are blocked, and \mathcal{A}_1 does not activate them.

The above lemmas conclude the proof of Theorem 2. When the total effect is identifiable, A_{γ} is an adjustment set, but it has a large size, so we provide a smaller adjustment set, defined as follows:

$$\mathcal{A}'_{\gamma} = \{ V_{t'} \in \mathcal{A}_{\gamma} | V \in Anc(X, \mathcal{G}^s) \cup Anc(Y, \mathcal{G}^s) \}.$$

Proposition 2. Consider an SCG \mathcal{G}^s and the total effect $P(y_t \mid do(x_{t-\gamma}))$, with $\gamma \geq 0$. Under assumptions of identifiability provided by Theorem 2, the set \mathcal{A}'_{γ} is an adjustment set for the total effect.

Proof. Let \mathcal{G}^f be an FTCG compatible with \mathcal{G}^s . Consider $V_{t'} \in \mathcal{A}_{\gamma} \backslash \mathcal{A}'_{\gamma}$: by definition of \mathcal{A}'_{γ} , it follows that $V_{t'} \not\in Anc(X_{t-\gamma}, \mathcal{G}^f) \cup Anc(Y_t, \mathcal{G}^f)$. Therefore $V_{t'}$ does not lie on any back-door path between $X_{t-\gamma}$

and Y_t : $V_{t'}$ is not necessary in the adjustment set, confirming that \mathcal{A}'_{γ} is also an adjustment set.

5 Conclusion

In this paper, we showed that the total effect is always identifiable for extended summary causal graphs (ECGs) but not always for summary causal graphs (SCGs). Additionally, we established necessary and sufficient graphical criteria for the identifiability of the total effect in SCGs, despite the potential presence of cycles.

These results have significant implications, such as impact analysis in dynamic systems, particularly in scenarios where experts are unable to provide either a full temporal causal graph (FTCG) or an ECG. They are also valuable in cases where the assumptions underlying causal discovery methods for inferring causal graphs with time lags are deemed overly restrictive.

We have provided a set of adjustment sets for estimating the total effect in ECGs, and an adjustment set when considering SCGs. However, those sets can potentially be large. While the estimation of the total effect based on such adjustments remains unbiased, there is a possibility of high variance. Determining the optimal adjustment set, that has minimal asymptotic estimation variance, remains an open question in this context.

References

Anand, T. V., Ribeiro, A. H., Tian, J., and Bareinboim, E. (2023). Causal effect identification in cluster dags.

Assaad, C. K., Devijver, E., and Gaussier, E. (2022a). Discovery of extended summary graphs in time series. In Cussens, J. and Zhang, K., editors, Proceedings of the Thirty-Eighth Conference on Uncertainty in Artificial Intelligence, volume 180 of Proceedings of Machine Learning Research, pages 96–106. PMLR.

Assaad, C. K., Devijver, E., and Gaussier, E.

- causal graphs in time series. Entropy, 24(8).
- Assaad, C. K., Devijver, E., and Gaussier, E. (2022c). Survey and evaluation of causal discovery methods for time series. J. Artif. Int. Res., 73.
- Assaad, C. K., Ez-Zejjari, I., and Zan, L. (2023). Root cause identification for collective anomalies in time series given an acyclic summary causal graph with loops. In Ruiz, F., Dy, J., and van de Meent, J.-W., editors, Proceedings of The 26th International Conference on Artificial Intelligence and Statistics, volume 206 of Proceedings of Machine Learning Research, pages 8395–8404. PMLR.
- Blondel, G., Arias, M., and Gavaldà, R. (2016). Identifiability and transportability in dynamic causal networks. International Journal of Data Science and Analytics, 3:131–147.
- Eichler, M. and Didelez, V. (2007). Causal reasoning in graphical time series models. In Proceedings of the Twenty-Third Conference on Uncertainty in Artificial Intelligence, UAI'07, page 109–116, Arlington, Virginia, USA. AUAI Press.
- Forré, P. and Mooij, J. M. (2017). Markov properties for graphical models with cycles and latent variables.
- Maathuis, M. and Colombo, D. (2013). A generalized backdoor criterion. The Annals of Statistics, 43.
- Meng, Y., Zhang, S., Sun, Y., Zhang, R., Hu, Z., Zhang, Y., Jia, C., Wang, Z., and Pei, D. (2020). Localizing failure root causes in a microservice through causality inference. In 2020 IEEE/ACM 28th International Symposium on Quality of Service (IWQoS), pages 1-10.
- Neyman, J., Dabrowska, D. M., and Speed, T. P. (1990). On the Application of Probability Theory to Agricultural Experiments. Essay on Principles. Section 9. Statistical Science, 5(4):465 – 472.
- Pearl, J. (1993). [Bayesian Analysis in Expert Systems]: Comment: Graphical Models, Causality and Intervention. Statistical Science, 8(3):266 – 269.

- Entropy-based discovery of summary Pearl, J. (1995). Causal diagrams for empirical research. Biometrika, 82(4):669-688.
 - Pearl, J. (2019). The seven tools of causal inference, with reflections on machine learning. Commun. ACM, 62(3):54–60.
 - Pearl, J. et al. (2000). Models, reasoning and inference. Cambridge, UK: CambridgeUniversityPress, 19:2.
 - Perkovic, E. (2020). Identifying causal effects in maximally oriented partially directed acyclic graphs. In Peters, J. and Sontag, D., editors, *Proceedings* of the 36th Conference on Uncertainty in Artificial Intelligence (UAI), volume 124 of Proceedings of Machine Learning Research, pages 530–539. PMLR.
 - Peters, J., Janzing, D., and Schölkopf, B. (2013). Causal inference on time series using restricted structural equation models. In Advances in Neural Information Processing Systems 26, pages 154–162.
 - Rubenstein, P. K., Bongers, S., Schölkopf, B., and Mooij, J. M. (2018). From deterministic ODEs to dynamic structural causal models. In Proceedings of the 34th Conference on Uncertainty in Artificial Intelligence (UAI), pages 114-123.
 - Shpitser, I. and Pearl, J. (2008). Complete identification methods for the causal hierarchy. Journal of Machine Learning Research, 9:1941–1979.
 - Shpitser, I., VanderWeele, T., and Robins, J. M. (2010). On the validity of covariate adjustment for estimating causal effects. In Proceedings of the Twenty-Sixth Conference on Uncertainty in Artificial Intelligence, UAI'10, page 527–536, Arlington, Virginia, USA. AUAI Press.
 - Spirtes, P., Glymour, C. N., Scheines, R., and Heckerman, D. (2000). Causation, prediction, and search. MIT press.