



HAL
open science

A non-asymptotic analysis of the single component PLS regression

Luca Castelli, Irène Gannaz, Clément Marteau

► **To cite this version:**

Luca Castelli, Irène Gannaz, Clément Marteau. A non-asymptotic analysis of the single component PLS regression. 2023. hal-04237950

HAL Id: hal-04237950

<https://hal.science/hal-04237950v1>

Preprint submitted on 11 Oct 2023

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

A non-asymptotic analysis of the single component PLS regression

Luca Castelli¹, Irène Gannaz², Clément Marteau¹

¹Univ Lyon, Université Claude Bernard Lyon 1, CNRS UMR 5208, Institut Camille Jordan, F-69622 Villeurbanne, France

²Univ. Grenoble Alpes, CNRS, Grenoble INP*, G-SCOP, 38000 Grenoble, France

October 11, 2023

Abstract

This paper investigates some theoretical properties of the Partial Least Square (PLS) method. We focus our attention on the single component case, that provides a useful framework to understand the underlying mechanism. We provide a non-asymptotic upper bound on the quadratic loss in prediction with high probability in a high dimensional regression context. The bound is attained thanks to a preliminary test on the first PLS component. In a second time, we extend these results to the sparse partial least squares (sPLS) approach. In particular, we exhibit upper bounds similar to those obtained with the lasso algorithm, up to an additional restricted eigenvalue constraint on the design matrix.

Keywords: *Partial least squares; dimension reduction; regression; sparsity*

1. Introduction

We are interested in the classical linear model in a high dimensional context. We observe a n -sample (X_i, Y_i) , $i = 1, \dots, n$, where the $Y_i \in \mathbb{R}$ are outcome variables of interest and the $X_i \in \mathbb{R}^p$ p -dimensional covariates. We consider a linear relationship within each couple (X_i, Y_i) , represented by the equation

$$Y = X\beta + \varepsilon, \tag{1}$$

where $\varepsilon = (\varepsilon_1, \dots, \varepsilon_n)^T \sim \mathcal{N}_n(0, \tau^2 I_n)$, $X = (X_1, \dots, X_n)^T \in \mathbb{R}^{n \times p}$ and $Y = (Y_1, \dots, Y_n)^T \in \mathbb{R}^n$. The matrix I_n is the identity matrix of size n and the parameter τ characterizes the noise level. The

*Institute of Engineering Univ. Grenoble Alpes

exponent T denotes the transpose operator. In this context, one might be alternatively interested in providing inference on the parameter β itself, or on $X\beta$ (prediction task). The regression model (1) has a long history. Several issues may arise, in particular in a high dimensional context, namely when p is of the same order, or much larger, than the number n of available observations. We refer to Giraud (2021) for a comprehensive introduction to this topic.

In this paper, we will focus our attention on the Partial Least Squares (PLS) principle. PLS was mainly developed in the chemometrics community (Martens and Naes, 1992). This approach has shown its ability for the prediction of regression models with many predictor variables (Garthwaite, 1994). It has been widely used in chemometrics (Wold, 1995; Wold, Sjöström, and Eriksson, 2001) but also in other fields such as social science (Sawatsky, Clyde, and Meek, 2015), and biology (Palermo, Piraino, and Zucht, 2009; Yang et al., 2017). Several extensions have been proposed over the years as, e.g., Delaigle and Hall (2012) for functional data or Naik and Tsai (2000) for the single-index models.

The idea of PLS is to seek a fixed number of directions – say $K \in \{1, \dots, p\}$ – formed by linear combinations of X coordinates, which are highly correlated with the target variable Y (see Mateos-Aparicio (2011) for a comprehensive introduction). These K directions are gathered in a weight matrix $W \in \mathbb{R}^{p \times K}$. The parameter β is then estimated by an appropriate linear combination of the columns of W . More formally, the PLS estimator satisfies

$$\hat{\beta}_W = \operatorname{argmin}_{w \in [W]} \|Y - Xw\|^2, \quad (2)$$

where $[W] \subset \mathbb{R}^p$ denotes the space spanned by the columns of the weight matrix W , and $\|\cdot\|$ is the ℓ^2 -norm on \mathbb{R}^n . The objective of the dimension reduction given by W is to decrease the number of features from p to K while retaining as much information as possible. On the contrary to Principal Components Reduction where the directions are built only considering the covariates X , PLS regression builds the weights W iteratively, considering the successive correlations with the outcome Y , to increase the prediction quality. We refer to Frank and Friedman (1993) and Krämer and Sugiyama (2011) and to Section 2 for more details regarding the way in which W is constructed.

Although the PLS principle has attracted a lot of attention over the years, few theoretical results have been obtained. Among others, we can mention Helland (1990) where the space $[W]$ resulting from the PLS approach has been characterised, or Cook, Li, and Chiaromonte (2010) and Cook, Helland, and Su (2013) for a connection between PLS regression and envelopes. While Chun and Keleş (2010) proved inconsistency of the PLS estimator when the number of covariates is too large, Cook and Forzani (2017) and Cook and Forzani (2019) established – up to strong constraints on β and the design matrix X – the asymptotic behavior of the mean squared error of prediction and prove that it may tend to 0 as the number of observations goes to infinity. Additional investigations have also proposed to take into account sparsity constraints in the PLS algorithm. We refer for instance to Durif et al. (2017) and Alsouki et al. (2023).

The aim of this paper is to investigate the theoretical performances of the PLS algorithm. We will focus our attention on the single component PLS algorithm, namely $K = 1$. Despite its apparent simplicity, this setting provides numerous statistical challenges. In particular, the non-linearity of

the corresponding estimator requires a careful attention. In this context, our aim is twofold:

- Obtaining non-asymptotic prediction bounds on the PLS estimator with high probability. In particular, we want to provide an extensive description of the performances and limitations of this approach with a minimal set of assumptions.
- Extend these results to a sparse scenario, and discuss existing similarities with alternative methods like the lasso estimator (see, e.g., Tibshirani (1996) and Section 2.3).

To this end, we establish in a first time a non-asymptotic bound on the prediction loss. Denoting by $\widehat{\beta}_{PLS}$ be the PLS estimator with $K = 1$, we prove in particular that with high probability,

$$\frac{1}{n} \|X\widehat{\beta}_{PLS} - X\beta\|^2 \leq B(\beta) + C\frac{\tau^2}{n} \max\left(\frac{\text{Tr}(\Sigma)}{\lambda}, \frac{\rho(\Sigma)\text{Tr}(\Sigma)}{\lambda^2}\right),$$

where C is a positive constant, $\text{Tr}(\cdot)$ is the trace operator on $\mathbb{R}^{p \times p}$, and $\Sigma = \frac{1}{n}X^TX \in \mathbb{R}^{p \times p}$ is the Gram matrix associated to the design X . The term λ corresponds to the norm of the first theoretical PLS component while $B(\beta)$ is a measure on the bias induced by the algorithm. The variance term can be read as a signal-to-noise ratio and allows to describe scenarios where the PLS estimator provides satisfying results. The formal result, along with an extended discussion, is displayed in Section 3.

Next, we extend this result to the parsimonious case using a sparse version $\widehat{\beta}_{sPLS}$ of the algorithm, including an ℓ_1 constraint in the optimisation process. Assuming that the Gram matrix Σ satisfies a restricted eigenvalue condition, we establish that, with high probability,

$$\frac{1}{n} \|X\widetilde{\beta}_{sPLS} - X\beta\|^2 \leq B(\beta) + C\frac{\tau^2 s}{n} \ln(p),$$

where s denotes the number of non-zero coefficient of the first PLS axis. In particular we recover, up to the bias term, the same kind of bound as those obtained for the Lasso procedure. See Section 4 for more details on our results and related consequences.

The contribution is structured as follows. Section 2 provides a brief overview on the PLS algorithm and on its extensions. We detail in particular the case of a projection on a single component, that is, $K = 1$. We also present the sparse algorithm studied in this paper. Section 3 gathers our first theoretical bound on the single PLS estimator and a discussion on some specific scenarios that may shed some light on the behaviour of the PLS principle. An extension to the sparse case is presented in Section 4, together with a corollary that involve an additional assumption on the Gram matrix Σ . The proofs of our main results are displayed in the Appendix.

All along the paper, we use the following notation. The ℓ^2 (resp. ℓ^1) norm in \mathbb{R}^p is written $\|\cdot\|_2$ (resp. $\|\cdot\|_1$), while $\|\cdot\|$ corresponds to the ℓ^2 norm in \mathbb{R}^n . For any matrix A , A^T corresponds to the transpose matrix, and $[A]$ to the vectorial space spanned by the columns of A . If A is a square matrix, we write $\text{Tr}(A)$ (resp. $\rho(A)$) for the trace (resp. spectral radius) of A . Here and below, the design matrix X is considered as deterministic. We denote by $\Sigma = X^TX/n$ the Gram matrix associated to the design X , $\widehat{\sigma} = X^TY/n$ the empirical covariance between X and the target Y and $\sigma = \mathbb{E}[\widehat{\sigma}] = \Sigma\beta$ its population counterpart.

2. A short PLS overview

In this section, we first recall the PLS algorithm. We will focus our attention on the single component case, and will briefly discuss existing results in this particular setting. Next, we present a sparse extension of this approach that includes a lasso type penalization in the optimisation process.

2.1. The PLS algorithm

For a fixed number of components $K \in \{1, \dots, p\}$, the PLS algorithm summarizes some directions w_k for $k \in \{1, \dots, K\}$ whose associated components t_k are the most possible correlated with the target variable Y . The K PLS components $(t_k)_{k=1, \dots, K}$ form an orthogonal basis of \mathbb{R}^p . They are build to compress the data in uncorrelated terms at each iterative step. The corresponding weights $W = (w_k)_{k=1, \dots, K}$ are computed iteratively, seeking the components which are the most correlated with the predicted variable Y on residuals.

For each new iteration, the algorithm advances by executing a deflation step: it takes the influence of the previously computed components out of the design matrix X . Let $X^{(1)} = X$ and for $k \in \{2, \dots, K\}$, define

$$X^{(k)} = X - P_{[t_1, t_2, \dots, t_{k-1}]}(X) = X^{(k-1)} - P_{[t_{k-1}]}(X^{(k-1)}),$$

where $P_{[t_1, t_2, \dots, t_k]}$ denotes the orthogonal projection on the subspace spanned by t_1, t_2, \dots, t_k . The weight w_k is then solution of the following optimisation problem

$$w_k = \underset{w \in \mathbb{R}^p}{\operatorname{argmin}}(-Y^T X^{(k)} w) \quad \text{s.t.} \quad \|w\|_2 = 1. \quad (3)$$

Components t_k are defined as $t_k = X^{(k)} w_k$, for all $k = 1, \dots, K$.

The formal PLS algorithm is displayed in [Algorithm 1](#) below.

Algorithm 1 PLS Algorithm

Input \mathbf{X}, Y and K

$\mathbf{X}_1 = \mathbf{X}$

for $k=1, \dots, K$ do

$\mathbf{w}_k = \mathbf{X}^{(k)T} Y / \|\mathbf{X}^{(k)T} Y\|_2$ (loadings computation)

$\mathbf{t}_k = \mathbf{X}^{(k)} \mathbf{w}_k$ (component construction)

$\mathbf{X}^{(k+1)} = \mathbf{X}^{(k)} - \mathbf{P}_{[\mathbf{t}_k]}(\mathbf{X}^{(k)})$ (deflation step)

end for

This algorithm is described for example in Helland (1990) where some of its properties are discussed. Alternative versions of this algorithm have been proposed over the years. For example, the components t_k , $k \in \{1, \dots, K\}$ can be calculated using variants as the NIPALS (non linear iterative least squares) introduced in Wold (1975) or SIMPLS (straightforward implementation of a statistically inspired modification of the PLS method) proposed by de Jong (1993) which com-

puts linear combinations of the original variables. We refer to Qin, Liu, and Tang (2023) for an overview of PLS algorithms, as well as the equivalence with the conjugate gradient method.

The PLS components (t_1, \dots, t_K) can be gathered in a matrix $T \in \mathbb{R}^{n \times K}$, where each column k of T corresponds to the component t_k . In particular, it can be noticed that $[T] = [XW]$ where $W = (w_1, \dots, w_K) \in \mathbb{R}^{p \times K}$ is the weight matrix. The PLS prediction $\hat{Y}_{PLS} := X\hat{\beta}_{PLS}$ is given by

$$\hat{Y}_{PLS} = T(T^T T)^{-1} T^T Y = P_{[XW]}(Y).$$

It immediately follows that the corresponding estimator of β is computed as

$$\hat{\beta}_{PLS} = \hat{\beta}_W = W(W^T \Sigma W)^{-1} W^T \hat{\sigma}, \quad (4)$$

where $\hat{\sigma} = X^T Y / n$ denotes the empirical covariance vector between X and Y .

The iterative nature of the PLS algorithm makes the corresponding estimation problem difficult to handle. However, Helland (1990) proved that $[W] = \hat{\mathcal{G}}$, where $\hat{\mathcal{G}}$ denotes the Krylov space which is the space generated by $\{\hat{\sigma}, \Sigma \hat{\sigma}, \dots, \Sigma^{K-1} \hat{\sigma}\}$. In particular, the optimisation problem (2) can be rewritten as

$$\hat{\beta}_{PLS} \in \underset{w \in \hat{\mathcal{G}}}{\operatorname{argmin}} \|Y - Xw\|^2. \quad (5)$$

The latter provides an alternative expression which is sometimes useful from a mathematical point of view (see for instance Cook and Forzani, 2019). Nevertheless, whatever the optimisation form, the resulting PLS estimator is always non-linear in the target Y . This induces several issues for obtaining accurate prediction bounds. In this context, we will focus in this paper on the single component case ($K = 1$). This case is a starting point preceding further additional investigations.

2.2. The single component PLS

In the specific case where the number of component K is constrained to be equal to 1, an explicit and workable expression of $\hat{\beta}_{PLS}$ can be obtained. Indeed, a direct application of (3), (4) or (5) leads to

$$\hat{\beta}_{PLS} = \frac{\hat{\sigma}^T \hat{\sigma}}{\hat{\sigma}^T \Sigma \hat{\sigma}} \hat{\sigma}. \quad (6)$$

In particular, it can be noticed that the estimator $\hat{\beta}_{PLS}$ is handled by the empirical covariance $\hat{\sigma}$ which corresponds to the first PLS direction w_1 up to a normalizing constant.

The formulation (6) has been at the core of the investigations presented in Cook and Forzani (2017). The latter established, up to our knowledge, the first consistency results for this estimator in several scenarios, including in particular high dimensional frameworks. Nevertheless, a central assumption made by the authors is the fact that the so-called population version of the Krylov space is of dimension 1. More formally, writing $\bar{\sigma} = \bar{\Sigma} \beta$ and $\bar{\Sigma} = \mathbb{E}(\Sigma)$ (assuming that the covariates are random variables), they suppose that $\dim \bar{\mathcal{G}} = 1$ where $\bar{\mathcal{G}}$ is the vectorial space generated by $\{\bar{\sigma}, \bar{\Sigma} \bar{\sigma}, \dots, \bar{\Sigma}^{p-1} \bar{\sigma}\}$. An immediate consequence of this assumption is that the coefficient vector β

is strongly constrained. Indeed

$$\dim(\bar{\mathcal{G}}) = 1 \quad \Rightarrow \quad \beta \in [\bar{\sigma}],$$

provided $\bar{\Sigma}$ is invertible. Moreover, $\dim(\bar{\mathcal{G}}) = 1$ implies that $\bar{\sigma}$ is an eigenvector of $\bar{\Sigma}$.

In [Section 3](#), we propose a bound on the prediction error with a controlled probability that does not use the assumption $\dim(\bar{\mathcal{G}}) = 1$. Moreover, our bound is non-asymptotic and does not require a specific regime for n and p . This non-asymptotic analysis shed lights on some particular scenarios, missing from [Cook and Forzani \(2017\)](#), where the term $\sigma^T \Sigma \sigma$ is too small to ensure a precise control of the denominator in [\(6\)](#).

2.3. Sparse-PLS

In presence of a high number of regressors, one might want to promote dimension reduction during the estimation process. Following [Tibshirani \(1996\)](#), a possible strategy is to keep the more relevant variables and to shrink the others to zero by adding a ℓ^1 -penalty on the coefficients in the objective function. Using this principle, the Sparse Partial Least Squares (sPLS) algorithm uses a sparse constraint to select a subset of variables that have the highest correlation with the response variable in the construction of the components. It is an effective tool for variable selection, dimension reduction, and developing predictive models using a limited set of relevant predictors. The sPLS estimator has been considered in several frameworks and application fields as, e.g., [Lee et al. \(2011\)](#) and [Abdel-Rahman et al. \(2014\)](#) in chemometrics, [Cao et al. \(2008\)](#) and [Chun and Keleş \(2010\)](#) in genomics or [Fuentes, Poncela, and Rodríguez \(2015\)](#) in econometrics.

Following [Durif et al. \(2017\)](#), a sparsity constraint is included in the estimation process by adding a ℓ^1 -penalty in the optimisation problem [\(3\)](#). More formally, the latter is replaced by

$$\tilde{w}_k = \underset{w \in \mathbb{R}^p}{\operatorname{argmin}} \left[-\frac{1}{n} Y^T X^{(k)} w + \mu \|w\|_1 \right] \quad \text{s.t.} \quad \|w\|_2 = 1, \quad (7)$$

where $\mu > 0$ denotes a regularization parameter that determines the sparsity level of the solution and where, for $k \in \{2, \dots, K\}$, $X^{(k)}$ is a deflated version of X on the previous components $\tilde{t}_1, \dots, \tilde{t}_{k-1}$, with $\tilde{t}_j = X^{(j)} \tilde{w}_j$, $j \leq k-1$. Several alternative approaches have been proposed over the last two decades to try to force sparsity of the solution. We refer, e.g. to [Cao et al. \(2008\)](#) and [Chun and Keleş \(2010\)](#).

[Durif et al. \(2017\)](#) provide a closed-form solution for the optimisation problem using proximal operator. This expression can be made even simpler in the specific case where $K = 1$ which is at the core of this paper. In particular, the corresponding estimator $\hat{\beta}_{sPLS}$ involves a soft-thresholded version of $\hat{\sigma}$ (see [Donoho and Johnstone \(1994\)](#) and [Proposition C.1](#)).

In [Section 4](#), we provide a complete non-asymptotic analysis of the prediction error associated with the estimator $\hat{\beta}_{sPLS}$, still in the single component case. We exhibit and discuss several scenarios where this algorithm provides accurate predictions in a high dimensional context. These scenarios include in particular the specific cases where the Gram matrix Σ satisfies a restricted eigenvalue

condition (see for instance Bickel, Ritov, and Tsybakov, 2009).

3. Theoretical results with single component PLS

Our first contribution is a non asymptotic control of the prediction loss, with an explicit upper bound. We also take into account in our procedure the reliability of the PLS algorithm. That is, we identify in the estimation scheme when the PLS regression does not bring enough information to ensure a good quality of prediction. Our estimator is modified accordingly.

3.1. Non-asymptotic control

We can notice from (6) that the single component PLS estimator is made of a direction $\hat{\sigma}$ and a so-called intensity $\hat{\sigma}^T \hat{\sigma} / \hat{\sigma}^T \Sigma \hat{\sigma}$. A random term appears in the denominator of this last quantity. If the deterministic counterpart of this random part is too small, in a sense which is made precise in the proof, we cannot expect to have an accurate control of the variance term. To this end, we slightly modify the PLS estimator and introduce a threshold for the denominator. Namely $\hat{\beta}_{PLS}$ is considered for the estimation if and only if $\hat{\sigma}^T \Sigma \hat{\sigma}$ is large enough compared to a minimal reference value p_n (see Theorem 3.1 for more detail). We stress that

$$\hat{\sigma}^T \Sigma \hat{\sigma} = \frac{1}{n} \|X \hat{\sigma}\|^2 = \frac{1}{n} \|\hat{t}_1\|^2,$$

where \hat{t}_1 denotes, up to a normalizing constant, the first PLS component. Intuitively, if the norm of the first component is close to zero, the PLS estimation will not be accurate and the estimator is replaced by 0. On the other hand, if its value is above a fixed level of inertia, we use the PLS regression algorithm. Applying this principle, which can be considered as a pre-processing testing procedure, we obtain the following theorem.

Theorem 3.1. *Let $\delta \in (0, 1)$. Define*

$$\hat{\beta}_\delta = \begin{cases} \hat{\beta}_{PLS} & \text{if } \hat{\sigma}^T \Sigma \hat{\sigma} > t_\delta p_n \\ 0 & \text{otherwise} \end{cases} \quad \text{with} \quad p_n = \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma),$$

for some explicit threshold constant t_δ depending only of δ . Then, with a probability higher than $1 - \delta$, there exist a constant $C_\delta > 0$, depending on δ , such that

$$\frac{1}{n} \|X \hat{\beta}_\delta - X \beta\|^2 \leq \frac{2}{n} \inf_{v \in [\sigma]} \|X(\beta - v)\|^2 + C_\delta \frac{\tau^2}{n} \max \left(\frac{\text{Tr}(\Sigma)}{\lambda}, \frac{\rho(\Sigma) \text{Tr}(\Sigma)}{\lambda^2} \right), \quad (8)$$

where

$$\lambda = \frac{\sigma^T \Sigma \sigma}{\|\sigma\|_2^2}. \quad (9)$$

The proof of Theorem 3.1 is postponed to Appendix B. It is mainly based on non-asymptotic deviation results on non-centered weighted χ^2 distributions. The proof manages the different values of the indicator function $\mathbb{1}_{\{\hat{\sigma}^T \Sigma \hat{\sigma} \geq t_\delta p_n\}}$ where t_δ is defined in (29). In particular, we prove that the

latter is equal to 1 when the amount of signal inside the first (theoretical) PLS component $\sigma^T \Sigma \sigma$ is large enough. On the other hand, we get an upper bound on the prediction loss when $\widehat{\beta}_\delta = 0$. This situation corresponds to the case where the first component does not contain significant information. We highlight the fact that it implies that $X\beta$ is relatively low, compared to the level of noise. The analysis of the bound displayed in [Theorem 3.1](#) is discussed in the next subsection.

3.2. Discussion

The bound displayed in [Theorem 3.1](#) is composed of two terms. The first one, equal to

$$\frac{2}{n} \inf_{v \in [\sigma]} \|X(\beta - v)\|^2,$$

represents a bias term. It measures the distance between the true signal $X\beta$ and the best possible prediction based on the first theoretical PLS axis $[\sigma]$. This term cancels out provided $\beta \in [\sigma]$. This situation occurs for instance in the case where both the dimension of the Krylov space $\mathcal{G} = \text{Vect} \{ \sigma, \Sigma \sigma, \dots, \Sigma^{p-1} \sigma \}$ is equal to 1 and the matrix Σ is invertible. This assumption was – in a slight different version – at the core of [Cook and Forzani \(2017\)](#).

The second term displayed in the right hand side of (8) can be considered as a variance term. It essentially measures the impact of the noise ε on the PLS algorithm. It involves a ratio between the trace of Σ and the term λ introduced in (9). The latter exactly corresponds to the norm of the first theoretical PLS component $t_1 := \frac{1}{\sqrt{n}} X w_1$ associated to the first normalized theoretical PLS axis $w_1 = \sigma / \|\sigma\|_2$. In other words, the term

$$1 \leq \frac{\text{Tr}(\Sigma)}{\lambda} \leq +\infty,$$

can be considered as an inverse relative inertia. This interpretation makes sense in the particular case where $\dim(\mathcal{G}) = 1$. In such a case, it can be proved that the first PLS axis w_1 is an eigenvector of Σ associated to the eigenvalue λ (see [Cook and Forzani, 2017](#)). In the general case, namely when $\dim(\mathcal{G})$ is not necessarily equal to 1, it provides in some sense an inverse of a signal-to-noise ratio that controls the accuracy of the single PLS component. If this ratio is close to 1, the first PLS component captures most of the inertia of the data and we obtain a variance term with a parametric rate τ^2/n . On the other hand, if this ratio is large, we cannot expect to get accurate results for the single component PLS: the amount of signal captured in the first component is not large enough to counterbalance the presence of the noise in the data.

This discussion can be illustrated by considering two extreme examples. First, assume that $\Sigma = I_p$, the identity matrix in $\mathbb{R}^{p \times p}$. In this particular case, we get that $\text{Tr}(\Sigma) = p$ and $\lambda = 1$. The variance term in [Theorem 3.1](#) is then of order $\tau^2 p/n$, which can be dramatically large in a high dimension setting. This appears to be quite natural since in such a situation, the explanatory variables are uncorrelated: nothing can be expected from the PLS algorithm which is designed to express the couple (Y, X) in a low dimensional space. On the other hand, the case where Σ has a rank equal to one with largest eigenvalue λ is the most favorable case since in such a setting, $\text{Tr}(\Sigma)/\lambda = 1$ and the variance term is equal to τ^2/n .

To conclude this discussion, we stress that our bound on the prediction loss is a generalization of the one displayed in Cook and Forzani (2017). It nevertheless differs from this former result by several items. First, as discussed above, it does not impose an assumption on the dimension of the Krylov space \mathcal{G} and on the rank of X . An immediate consequence is the presence of a bias term. Secondly, our bound is non-asymptotic: we provide a control with high probability, whatever the values of n and p . Such an approach allows to consider all the possible values of $\|w_1\|_2$ and $\|t_1\|$: we do not require that these quantities are bounded from below. In particular, the fact the our variance term is slightly different from Cook and Forzani (2017) is an immediate consequence of this approach.

4. Sparse estimation

4.1. Statistical performances of the single component s-PLS estimator

First, we focus our attention on the sparse version of the single component PLS estimator. In particular, we consider the estimator $\hat{\beta}_{sPLS}$ defined in (4) with $W = \tilde{w}_1$ where \tilde{w}_1 is solution of the optimisation problem (7). Recall that the latter aims at providing a sparse approximation of the first theoretical PLS axis $w_1 = \sigma/\|\sigma\|_2$. We are hence implicitly interested in the situations where the number of non-zero coefficients of $\sigma \in \mathbb{R}^p$ is small compared to the number of available data n . In particular, we take advantage of the expression of $\hat{\beta}_{sPLS}$ which involves a shrinkage operator.

Denote by $\mathcal{J}_0 = \{j = 1, \dots, p, \sigma_j \neq 0\}$ the support of σ . Hereafter, a vector $v_{\mathcal{I}}$ denotes the vector v where all v_j are set to 0 for $j \notin \mathcal{I}$. The purpose of this section is to achieve bounds on the quadratic loss in prediction with terms depending only on the variables which contribute to the first component. Writing $\Sigma_{\mathcal{J}_0} = \frac{1}{n} X_{\mathcal{J}_0}^T X_{\mathcal{J}_0}$ with $(X_{\mathcal{J}_0})_{ij} = 0$ if $j \notin \mathcal{J}_0$, $i = 1, \dots, p$, we expect to reduce the variance appearing in (8) by exhibiting terms of order $\text{Tr}(\Sigma_{\mathcal{J}_0})$ instead of $\text{Tr}(\Sigma)$. Before presenting our results, we need two different assumptions on the model.

Assumption A.1. The columns of X are normalized, namely

$$\Sigma_{jj} = \frac{1}{n} \sum_{i=1}^p X_{ij}^2 = 1 \quad \forall j \in \{1, \dots, p\}.$$

This assumption is quite standard for the linear model (1). It allows to get simpler expression for the penalty level μ that appears in the optimisation process (7). This assumption can be guaranteed thanks to a scaling of the explanatory variables.

Assumption A.2. Let $\delta \in (0, 1)$ be fixed. There exists a term $d_{\delta,p}$ such that

$$\sigma^T \Sigma \sigma > d_{\delta,p} \frac{\tau^2}{n} \rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0}).$$

We recall that, up to a normalizing constant, the quantity $\sigma^T \Sigma \sigma$ corresponds to the norm of the first theoretical PLS component. [Assumption A.2](#) can hence be understood as minimal energy on the signal. In [Section 3](#), we have discussed the fact that situations where $\sigma^T \Sigma \sigma$ is small create numerous issues in the control of the prediction loss. These issues have been circumvented in [Theorem 3.1](#) thanks to the introduction of a threshold on the denominator. In a sparse context, constructing a test allowing to get rid of this assumption appears to be quite involved. It would require in particular a precise knowledge on the location of the support \mathcal{J}_0 . This is not reasonable in practice. Nevertheless, we will present in [Section 4.2](#) below an alternative procedure that allow to remove [Assumption A.2](#), up to an additional constraint on the design matrix.

A control for the prediction loss associated to the sPLS estimator $\widehat{\beta}_{sPLS}$ is displayed below. The proof is postponed to [Appendix D.1](#).

Theorem 4.1. *Let $\delta \in (0, 1/2)$ be fixed. Assume that [Assumption A.1](#) and [Assumption A.2](#) are satisfied with $d_{\delta,p} = C_0 \left(\ln\left(\frac{10}{\delta}\right) + \ln\left(\frac{p}{\delta}\right) \right)$, with $C_0 > 0$ an explicit constant. Let*

$$\mu = 2\tau \sqrt{\frac{2}{n} \ln\left(\frac{2p}{\delta}\right)}. \quad (10)$$

Then, we get, with probability greater than $1 - \delta$,

$$\frac{1}{n} \|X(\widehat{\beta}_{sPLS} - \beta)\|^2 \leq \frac{2}{n} \inf_{v \in [\sigma]} \|X(\beta - v)\|^2 + D_\delta \frac{\tau^2 s}{n} \max\left(\frac{\rho(\Sigma_{\mathcal{J}_0})}{\lambda^2}, \frac{1}{\lambda}\right) \ln\left(\frac{p}{\delta}\right), \quad (11)$$

where $s := |\mathcal{J}_0| = \text{Tr}(\Sigma_{\mathcal{J}_0})$ denotes the size of the support and D_δ a positive constant depending only on δ .

The constant C_0 is not detailed here for the sake of clarity. It is given explicitly in the proof, in [Appendix D.1.4](#).

The variance term in the bound obtained above differs from the one displayed in [Theorem 3.1](#) (although the bias remains the same). This an immediate consequence of the ℓ_1 constraint introduced during the optimisation process. While the first PLS direction \widehat{w}_1 is equal to $\widehat{\sigma}$ (up to a normalizing constant) in the framework of [Section 3](#), here we have $\widehat{w}_1 = \widetilde{\sigma}$ where $\widetilde{\sigma}$ is a thresholded version of $\widehat{\sigma}$. In particular, the support of $\widetilde{\sigma}$ is expected to be close (in a sense which is made precise in the proof) of \mathcal{J}_0 . As a consequence, the variance associated to the estimation of the σ is

$$\frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}) = \frac{\tau^2 s}{n} \leq \frac{\tau^2}{n} \text{Tr}(\Sigma),$$

where for the first equality, we have used [Assumption A.1](#). In the same spirit, we can notice that $\rho(\Sigma_{\mathcal{J}_0}) \leq \rho(\Sigma)$. In particular, we can expect a significant improvement of our bounds in the situation where the first theoretical PLS direction $w_1 = \sigma$ is sparse, namely when $|\mathcal{J}_0| = s \ll p$. The counterpart of this improvement is a term of order $\log(p)$ in the bound. Such a term is quite standard in the literature.

As for the standard framework ([Section 3](#)), the ratio between the amount of signal available in the first PLS component described by λ and the noise term discussed above plays an important role in the behaviour of the method. We can get rid of this ratio by using a classical assumption when

dealing with sparsity constraints.

Assumption A.3 (Restricted eigenvalue condition) There exists a constant $\phi > 0$ such that,

$$\min_{\gamma: \|\gamma_{\mathcal{J}_0^c}\|_1 \leq 3\|\gamma_{\mathcal{J}_0}\|_1} \frac{1}{n} \frac{\|X\gamma\|^2}{\|\gamma\|_2^2} \geq \frac{1}{\phi}.$$

This kind of assumption has been at the core of several contributions and discussions. It requires that for any $\gamma \in \mathbb{R}^p$, the norm of $\|X\gamma\|/\sqrt{n}$ is comparable to $\|\gamma\|_2$ provided the signal in γ is mainly concentrated in \mathcal{J}_0 . We refer to Bickel, Ritov, and Tsybakov (2009) among others, where this assumption is discussed and compared to other types of constraints. In the following, we apply this assumption on some vectors $\gamma \in \mathbb{R}^p$ whose support is included in \mathcal{J}_0 . Accordingly, [Assumption A.3](#) could be weakened by just assuming that the restricted matrix $X_{\mathcal{J}_0}$ is full rank. The following result is an almost immediate consequence of [Theorem 4.1](#) that takes advantage of [Assumption A.3](#).

Corollary 4.2. *Let $\delta \in (0, 1/2)$ be fixed. Assume that [Assumption A.1](#), [Assumption A.2](#) and [Assumption A.3](#) are satisfied. Setting*

$$\mu = 2\tau \sqrt{\frac{2}{n} \ln \left(\frac{2p}{\delta} \right)},$$

we get, with probability greater than $1 - \delta$,

$$\frac{1}{n} \|X(\hat{\beta}_{sPLS} - \beta)\|^2 \leq \frac{2}{n} \inf_{v \in [\sigma]} \|X(\beta - v)\|^2 + D_\delta \frac{\tau^2 s}{n} \max(\rho(\Sigma_{\mathcal{J}_0}), 1) \ln \left(\frac{p}{\delta} \right), \quad (12)$$

for some positive constant D_δ .

The proof is postponed to [Appendix D.2](#). [Assumption A.3](#) allows to lower bound the quantity λ . Up to the spectral radius $\rho(\Sigma_{\mathcal{J}_0})$ (which can reasonably be assumed to be bounded), we hence obtain a variance term

$$D_\delta \frac{\tau^2}{n} \ln \left(\frac{p}{\delta} \right),$$

which is exactly of same order as those obtained for the Lasso estimator and its variants. We refer, e.g., to Bickel, Ritov, and Tsybakov (2009), Meinshausen (2013) or Dalalyan, Hebiri, and Lederer (2017).

4.2. An alternative procedure

We recall that the results displayed in [Theorem 4.1](#) heavily rely on [Assumption A.2](#) which in some sense requires a minimal energy in the first theoretical PLS component. To get round of this issue, we slightly modify the single component sparse PLS estimator $\hat{\beta}_{sPLS}$. Recall the latter is defined in [\(4\)](#) with $W = \tilde{w}_1$ and where \tilde{w}_1 is solution of the optimisation problem [\(7\)](#). Using simple algebra, we can in particular establish that

$$\hat{\beta}_{sPLS} = (\tilde{w}_1^T \Sigma \tilde{w}_1)^{-1} \tilde{w}_1^T \hat{\sigma} \times \tilde{w}_1, \quad (13)$$

(see [Proposition C.1](#)). Under [Assumption A.2](#), the inverse of $\tilde{w}_1^T \Sigma \tilde{w}_1$ exists with high probability. We propose here to replace $\hat{\sigma}$ in (13) by a thresholded version $\tilde{\sigma}$,

$$\tilde{\sigma}_j = \text{sgn}(\hat{\sigma}_j)(|\hat{\sigma}_j| - \mu)_+ \quad \forall j \in \{1, \dots, p\},$$

where for any $x \in \mathbb{R}$, $\text{sgn}(x) = 1$ if $x > 0$, -1 if $x < 0$, and 0 if $x = 0$, and where $x_+ = x$ if $x \geq 0$ and 0 otherwise. This heuristic leads to the the following estimator

$$\tilde{\beta} = (\tilde{w}_1^T \Sigma \tilde{w}_1)^{-1} \tilde{w}_1^T \tilde{\sigma} \times \tilde{w}_1 := \tilde{\lambda}^{-1} \tilde{w}_1. \quad (14)$$

By taking advantage of the restricted eigenvalue condition ([Assumption A.3](#)), we are able to manage the values of $\tilde{\lambda}$ without using [Assumption A.2](#). This discussion is formalized in the following theorem.

Theorem 4.3. *Let $\delta \in (0, 1/2)$ be fixed and $\tilde{\beta}$ the estimator introduced in (13). Assume that [Assumption A.1](#) and [Assumption A.3](#) are satisfied. Setting*

$$\mu = 2\tau \sqrt{\frac{2}{n} \ln \left(\frac{p}{\delta} \right)},$$

we get, with probability greater than $1 - \delta$,

$$\frac{1}{n} \|X(\tilde{\beta} - \beta)\|^2 \leq \frac{2}{n} \inf_{v \in [\sigma]} \|X(\beta - v)\|^2 + D'_\delta \frac{\tau^2 s}{n} \max(\rho(\Sigma_{\mathcal{J}_0}), 1) \ln \left(\frac{p}{\delta} \right), \quad (15)$$

where $s := |\mathcal{J}_0| = \text{Tr}(\Sigma_{\mathcal{J}_0})$ denotes the size of the support and D'_δ a positive constant depending only on δ .

The proof is displayed in [Appendix D.3](#). The bound displayed in [Theorem 4.3](#) can be compared to the one presented in [Corollary 4.2](#). We obtain, up to some constants, exactly the same bounds, but without requiring [Assumption A.2](#).

5. Conclusion

In this contribution, we have provided non-asymptotic bounds on the prediction loss for the single component PLS estimator, including a sparse version. Our aim was to introduce minimal assumptions on the model allowing a control on the prediction loss. This allows in particular to shed light on several scenarios where the PLS estimator can lead to interesting and pertinent predictions.

The different contributions displayed all along our paper ([Theorems 3.1, 4.1 and 4.3](#) in particular) entail that the PLS algorithm does not immediately lead to satisfying theoretical results. Indeed, we need either strong additional constraints on the signal ([Assumption A.2](#)) or the introduction of a pre-processing step in the algorithm: a test ([Theorem 3.1](#)) or a soft-thresholding rule for $\hat{\sigma}$ ([Theorem 4.3](#)).

Our investigations are limited to the single component PLS estimator. We can indeed notice that this setting already brings a lot of technical issues. Nevertheless, it is expected to propose a similar

non-asymptotic treatment in the general case with K PLS components for some $K \in \{1, \dots, p\}$. This might be the core of a future contribution.

Acknowledgements

The authors would like to dedicate this contribution to François Wahl who suddenly passed away in April 2022. He initiated discussions on the PLS algorithm some years ago that created new research opportunities in our team.

A. Technical results

This section is dedicated to some specific technical results that will be used all along the proofs.

A.1. Distributions of $\hat{\sigma}$ and $\Sigma^{\frac{1}{2}}\hat{\sigma}$

We first state the moments and the distribution of the main quantities appearing in the construction of the single component PLS estimator.

Lemma A.1. *We have*

$$\hat{\sigma} \sim \mathcal{N}_p\left(\sigma, \frac{\tau^2}{n}\Sigma\right) \quad \text{and} \quad \Sigma^{\frac{1}{2}}\hat{\sigma} \sim \mathcal{N}_p\left(\Sigma^{\frac{1}{2}}\sigma, \frac{\tau^2}{n}\Sigma^2\right).$$

In particular

$$\mathbb{E}[\hat{\sigma}^T\hat{\sigma}] = \sigma^T\sigma + \frac{\tau^2}{n}\text{Tr}(\Sigma), \quad \mathbb{E}[\hat{\sigma}^T\Sigma\hat{\sigma}] = \sigma^T\Sigma\sigma + \frac{\tau^2}{n}\text{Tr}(\Sigma^2),$$

and

$$\mathbb{E}[(\hat{\sigma} - \sigma)^T\Sigma(\hat{\sigma} - \sigma)] = \frac{\tau^2}{n}\text{Tr}(\Sigma^2).$$

The results of this lemma are a direct consequence of Model (1) and of the fact that $\varepsilon \sim \mathcal{N}(0, \tau^2 I_n)$. The proof is thus omitted.

A.2. Deviation inequalities on $\hat{\sigma}$ and $\Sigma^{1/2}\hat{\sigma}$

The following proposition is an extension of a result established in Laurent, Loubes, and Marteau (2012). It provides deviation inequalities for some quadratic function of a Gaussian vector.

Proposition A.2. *Let $U \sim \mathcal{N}_D(m, tA)$ with $D \in \mathbb{N}$, $m \in \mathbb{R}^D$, $t \in \mathbb{R}_+$ and $A \in \mathbb{R}^{D \times D}$ a symmetric positive matrix. Define, for any $s \in \mathbb{N}$,*

$$\Theta_s = t^2 \text{Tr}(A^{2(s+1)}) + 2t\rho(A^{s+1})\|A^{\frac{s}{2}}m\|^2,$$

Then, for all $s \in \mathbb{N}$ and $x \geq 0$,

$$\begin{aligned} i) \quad & \mathbb{P}\left(U^T A^s U - \mathbb{E}[U^T A^s U] \geq 2\sqrt{\Theta_s x} + 2t\rho(A)^{s+1}x\right) \leq e^{-x}, \\ ii) \quad & \mathbb{P}\left(U^T A^s U - \mathbb{E}[U^T A^s U] \leq -2\sqrt{\Theta_s x}\right) \leq e^{-x}. \end{aligned}$$

Proof. Denote by $(\lambda_i)_{i=1}^D$ the eigenvalues of A and $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_D)$. Set $\theta = P^T A^{\frac{s}{2}} m$, where P is a matrix verifying $A = P\Lambda P^T$. First remark that

$$U^T A^s U = U^T A^{\frac{s}{2}} P P^T A^{\frac{s}{2}} U = \|P^T A^{\frac{s}{2}} U\|^2,$$

Moreover, $P^T A^{\frac{s}{2}} U \sim \mathcal{N}(\theta, t\Lambda^{s+1})$ where Λ^{s+1} is a diagonal matrix. The result then follows from a direct application of Lemma 2 from Laurent, Loubes, and Marteau (2012) on the Gaussian vector $P^T A^{\frac{s}{2}} U$ and from the bound

$$\sum_{j=1}^D \lambda_j^{s+1} \theta_j^2 \leq \max_{j=1 \dots D} \lambda_j^{s+1} \times \|\theta\|^2 = \rho(A^{s+1}) \|A^{\frac{s}{2}} m\|^2.$$

□

Before stating additional results, we introduce, for any $x \in \mathbb{R}^+$, the following quantities:

$$\mathbf{T}_1(x) = g(x) \frac{\tau^2}{n} \text{Tr}(\Sigma) + 2\sqrt{2} \sqrt{\frac{\tau^2}{n}} \rho(\Sigma)^{\frac{1}{2}} \sqrt{x} \|\sigma\|_2, \quad (16)$$

$$\mathbf{T}_2(x) = g(x) \frac{\tau^2}{n} \text{Tr}(\Sigma^2) + 2\sqrt{2} \sqrt{\frac{\tau^2}{n}} \rho(\Sigma) \sqrt{x} \|\Sigma^{\frac{1}{2}} \sigma\|_2, \quad (17)$$

$$\mathbf{T}_3(x) = g(x) \frac{\tau^2}{n} \text{Tr}(\Sigma^2), \quad (18)$$

with

$$g(x) = 1 + 2x + 2\sqrt{x}. \quad (19)$$

The following proposition will be the core of the proof of our main results. It provides deviation results on the main quantities of interest.

Proposition A.3. For any $0 < \delta < 1$, let $(\mathcal{A}_{i,\delta})_{i=1}^3$ the events respectively defined as

$$\mathcal{A}_{1,\delta} = \{|\widehat{\sigma}^T \widehat{\sigma} - \sigma^T \sigma| \leq \mathbf{T}_1(x_\delta)\}, \quad (20)$$

$$\mathcal{A}_{2,\delta} = \{|\widehat{\sigma}^T \Sigma \widehat{\sigma} - \sigma^T \Sigma \sigma| \leq \mathbf{T}_2(x_\delta)\}, \quad (21)$$

$$\text{and } \mathcal{A}_{3,\delta} = \{(\widehat{\sigma} - \sigma)^T \Sigma (\widehat{\sigma} - \sigma) \leq \mathbf{T}_3(x_\delta)\}, \quad (22)$$

with $x_\delta = \ln(5/\delta)$. Then,

$$\mathbb{P}(\mathcal{A}_\delta) \geq 1 - \delta \quad \text{where} \quad \mathcal{A}_\delta := \mathcal{A}_{1,\delta} \cap \mathcal{A}_{2,\delta} \cap \mathcal{A}_{3,\delta}.$$

Remark. The bounds displayed in Lemma A.2 are sharp in some sense (up to the constant terms) and are particularly useful for testing issues. Since we are working in an estimation context, we do

not need such a level of accuracy. The quantities $\mathbf{T}_1(x_\delta)$, $\mathbf{T}_2(x_\delta)$ and $\mathbf{T}_3(x_\delta)$ respectively defined in (20), (21), (22) hence correspond to rough bounds on the deviation terms. In particular, in the proof below, we often use the inequality $\text{Tr}(\Sigma^2) \leq \text{Tr}^2(\Sigma)$ which is not sharp but well suited for our analysis.

Proof. First, applying item i) of Lemma A.2 on the variable $\hat{\sigma}$ with $s = 0$, $t = \frac{\tau^2}{n}$, $m = \sigma$ and $A = \Sigma$, we get

$$\mathbb{P}\left(\hat{\sigma}^T \hat{\sigma} \geq \sigma^T \sigma + \frac{\tau^2}{n} \text{Tr}(\Sigma) + 2\sqrt{x_\delta} \sqrt{\frac{\tau^4}{n^2} \text{Tr}(\Sigma^2) + 2\frac{\tau^2}{n} \rho(\Sigma) \|\sigma\|_2^2} + 2\frac{\tau^2}{n} \rho(\Sigma) x_\delta\right) \leq \frac{\delta}{5}.$$

Then, remark that

$$\frac{\tau^2}{n} \text{Tr}(\Sigma) + 2\sqrt{x_\delta} \sqrt{\frac{\tau^4}{n^2} \text{Tr}(\Sigma^2) + 2\frac{\tau^2}{n} \rho(\Sigma) \|\sigma\|_2^2} + 2\frac{\tau^2}{n} \rho(\Sigma) x_\delta \leq \mathbf{T}_1(x_\delta),$$

where we have used the bounds $\sqrt{\text{Tr}(\Sigma^2)} \leq \text{Tr}(\Sigma)$, $\rho(\Sigma) \leq \text{Tr}(\Sigma)$ and $\sqrt{a+b} \leq \sqrt{a} + \sqrt{b}$ for any $a, b \in \mathbb{R}_+$. Similarly,

$$\mathbb{P}\left(\hat{\sigma}^T \hat{\sigma} \leq \sigma^T \sigma + \frac{\tau^2}{n} \text{Tr}(\Sigma) - 2\sqrt{x_\delta} \sqrt{\frac{\tau^4}{n^2} \text{Tr}(\Sigma^2) + 2\frac{\tau^2}{n} \rho(\Sigma) \|\sigma\|_2^2}\right) \leq \frac{\delta}{5}.$$

Using again a rough bound

$$-\frac{\tau^2}{n} \text{Tr}(\Sigma) + 2\sqrt{x_\delta} \sqrt{\frac{\tau^4}{n^2} \text{Tr}(\Sigma^2) + 2\frac{\tau^2}{n} \rho(\Sigma) \|\sigma\|_2^2} \leq \mathbf{T}_1(x_\delta),$$

we get

$$\mathbb{P}(\mathcal{A}_{1,\delta}^c) \leq \frac{2\delta}{5}.$$

Using again Lemma A.2 with $s = 1$, $t = \frac{\tau^2}{n}$, $A = \Sigma$ and $m = \sigma$ (resp. $m = 0$), we obtain respectively

$$\mathbb{P}(\mathcal{A}_{2,\delta}^c) \leq \frac{2\delta}{5} \quad \text{and} \quad \mathbb{P}(\mathcal{A}_{3,\delta}^c) \leq \frac{\delta}{5}.$$

Using the union bound

$$\mathbb{P}(\mathcal{A}^c) \leq \mathbb{P}(\mathcal{A}_{1,\delta}^c) + \mathbb{P}(\mathcal{A}_{2,\delta}^c) + \mathbb{P}(\mathcal{A}_{3,\delta}^c),$$

allows to conclude the proof. \square

A.3. Control of $\hat{\lambda}^{-1}$

Recall that the single component PLS estimator of β can be written as

$$\hat{\beta} = \hat{\lambda}^{-1} \hat{\sigma} \quad \text{with} \quad \hat{\lambda}^{-1} = \frac{\hat{\sigma}^T \hat{\sigma}}{\hat{\sigma}^T \Sigma \hat{\sigma}}.$$

To get a bound for the prediction loss, we need to control the deviation of $\widehat{\lambda}^{-1}$ with respect to its deterministic counterpart

$$\lambda^{-1} = \frac{\sigma^T \sigma}{\sigma^T \Sigma \sigma}.$$

The stochastic term $\widehat{\sigma}^T \Sigma \widehat{\sigma}$ in the denominator creates statistical issues. To get rid of them, we need a minimal value for $\sigma^T \Sigma \sigma$ as presented in the following proposition.

Lemma A.4. *Assume that*

$$r \sigma^T \Sigma \sigma \geq g(x_\delta) \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma), \quad (23)$$

for some $r \in (0, 1)$. Then, on the event \mathcal{A}_δ defined in Proposition A.3, we have

$$\widehat{\lambda}^{-1} \lambda \leq C_{\delta, r},$$

for some explicit constant $C_{\delta, r}$ depending only on δ and r .

Proof. First write

$$\widehat{\lambda}^{-1} \lambda = \underbrace{\frac{\widehat{\sigma}^T \widehat{\sigma}}{\sigma^T \sigma}}_{:=S_1} \times \underbrace{\frac{\sigma^T \Sigma \sigma}{\widehat{\sigma}^T \Sigma \widehat{\sigma}}}_{:=S_2}.$$

We first concentrate our attention on S_2 . For any $s \in (0, 1)$,

$$\begin{aligned} -2\sigma^T \Sigma (\sigma - \widehat{\sigma}) &\geq -2\sqrt{\sigma^T \Sigma \sigma} \sqrt{(\widehat{\sigma} - \sigma)^T \Sigma (\widehat{\sigma} - \sigma)}, \\ &\geq -s\sigma^T \Sigma \sigma - s^{-1}(\widehat{\sigma} - \sigma)^T \Sigma (\widehat{\sigma} - \sigma), \end{aligned}$$

since $2ab \leq sa^2 + s^{-1}b^2$ for all $a, b > 0$. Then, on the event \mathcal{A}_δ introduced in Proposition A.3,

$$\begin{aligned} \widehat{\sigma}^T \Sigma \widehat{\sigma} &= (\widehat{\sigma} - \sigma)^T \Sigma (\widehat{\sigma} - \sigma) - 2\sigma^T \Sigma (\sigma - \widehat{\sigma}) + \sigma^T \Sigma \sigma, \\ &\geq (\widehat{\sigma} - \sigma)^T \Sigma (\widehat{\sigma} - \sigma) - s\sigma^T \Sigma \sigma - s^{-1}(\widehat{\sigma} - \sigma)^T \Sigma (\widehat{\sigma} - \sigma) + \sigma^T \Sigma \sigma, \\ &\geq (1-s)\sigma^T \Sigma \sigma + (1-s^{-1})(\widehat{\sigma} - \sigma)^T \Sigma (\widehat{\sigma} - \sigma), \\ &\geq (1-s)\sigma^T \Sigma \sigma - (s^{-1}-1)\mathbf{T}_3(x_\delta), \end{aligned}$$

where for the last two lines, we have used the fact that $s \in (0, 1)$. Provided (23) holds, we obtain, since $\text{Tr}(\Sigma^2) \leq \rho(\Sigma)\text{Tr}(\Sigma)$,

$$\widehat{\sigma}^T \Sigma \widehat{\sigma} \geq (1-s)\sigma^T \Sigma \sigma - (s^{-1}-1)r\sigma^T \Sigma \sigma = (1+r-s-s^{-1}r)\sigma^T \Sigma \sigma.$$

In order to obtain a positive bound, parameters s and r have to satisfy $(1-s) > (s^{-1}-1)r$, which holds as soon as $r < s$. Setting for instance $s = \frac{1+r}{2}$, we get

$$\frac{\sigma^T \Sigma \sigma}{\widehat{\sigma}^T \Sigma \widehat{\sigma}} \leq \frac{1}{1+r - \left(\frac{1+r}{2} + \frac{2r}{1+r}\right)}. \quad (24)$$

Now, we provide a bound on the term S_1 . Still on the event \mathcal{A}_δ , we have

$$\frac{\widehat{\sigma}^T \widehat{\sigma}}{\sigma^T \sigma} \leq \frac{\sigma^T \sigma + \mathbf{T}_1(x_\delta)}{\sigma^T \sigma} = 1 + \frac{\mathbf{T}_1(x_\delta)}{\sigma^T \sigma}.$$

Using the simple bound $\sigma^T \Sigma \sigma \leq \sigma^T \sigma \rho(\Sigma)$, we get that the constraint (23) induces the inequality

$$r\sigma^T \sigma \geq g(x_\delta) \frac{\tau^2}{n} \text{Tr}(\Sigma).$$

Using the definition of $\mathbf{T}_1(x_\delta)$, we hence obtain

$$\frac{\widehat{\sigma}^T \widehat{\sigma}}{\sigma^T \sigma} \leq 1 + r + 2\sqrt{2r \frac{x_\delta}{g(x_\delta)}}. \quad (25)$$

The conclusion follows from (24) and (25) with

$$C_{\delta,r} = \frac{1 + r + 2\sqrt{2r \frac{x_\delta}{g(x_\delta)}}}{1 + r - \left(\frac{1+r}{2} + \frac{2r}{1+r}\right)}. \quad (26)$$

□

B. Proof of Theorem 3.1

Introduce the parameter $\bar{\beta}$ as

$$\bar{\beta} = \frac{\sigma^T \sigma}{\sigma^T \Sigma \sigma} \sigma := \lambda^{-1} \sigma = \underset{u \in [\sigma]}{\text{argmin}} \|X\beta - Xu\|^2. \quad (27)$$

Using a simple bound, we can first notice that

$$\frac{1}{n} \|X(\widehat{\beta}_\delta - \beta)\|^2 \leq \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + \frac{2}{n} \|X(\widehat{\beta}_\delta - \bar{\beta})\|^2. \quad (28)$$

The first term in the r.h.s. of (28) exactly corresponds to the bias term appearing in Theorem 3.1.

Let

$$\widehat{\beta}_{\delta,r} = \frac{\widehat{\sigma}^T \widehat{\sigma}}{\widehat{\sigma}^T \Sigma \widehat{\sigma}} \widehat{\sigma} \times \Psi_{\delta,r} \quad \text{with} \quad \Psi_{\delta,r} = \mathbb{1}_{\{\widehat{\sigma}^T \Sigma \widehat{\sigma} > t_{\delta,r} p_n\}},$$

where $t_{\delta,r}$ is a threshold defined below in (29) and r is some parameter. We choose to keep this parameter here, to highlight the test-like procedure. In particular, r determines the low signal set (see below). Actually, Theorem 3.1 will be provided by taking $r = 1/2$. That is, with $\widehat{\beta}_\delta = \widehat{\beta}_{\delta,1/2}$ and $t_\delta = t_{\delta,1/2}$.

Until the end of the proof, we will focus our attention on the second term in the r.h.s. of (28). To this end, we consider three different scenarios:

- i) The low signal case: the norm $\|\sigma\|_2$ is small (in a sense which is made precise latter on). In such a case, our estimator $\widehat{\beta}_{\delta,r}$ is equal to 0 with high probability (w.h.p.).
- ii) The high signal case: the norm of σ exceeds a given level. Then, the indicator function in $\widehat{\beta}_{\delta,r}$ is equal to 1 w.h.p. and we use deviation results established in the previous section to obtain a bound in prediction.
- iii) The intermediate case: we are not able to control w.h.p. the behaviour of the indicator

function, but we take advantage of the bounds obtained in both previous cases.

B.1. Low signal case (Case i)

We introduce $\mathcal{S}_L(C)$ where, for any $C \in \mathbb{R}^+$,

$$\mathcal{S}_L(C) = \{\sigma \in \mathbb{R}^p, \sigma^T \Sigma \sigma \leq Cp_n\} \quad \text{and} \quad p_n = \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma).$$

We assume in this section that $\sigma \in \mathcal{S}_L(g(x_\delta)/r)$ where $r \in (0, 1)$ appears in [Lemma A.4](#). Firstly, we prove that in this case $\Psi_{\delta,r} = 0$ on the event \mathcal{A}_δ . Indeed, according to (17) and (21), we have

$$\begin{aligned} \hat{\sigma}^T \Sigma \hat{\sigma} &\leq 2\sigma^T \Sigma \sigma + g(x_\delta) \frac{\tau^2}{n} \text{Tr}(\Sigma^2) + 2\frac{\tau^2}{n} \rho(\Sigma)^2 x_\delta \\ &\leq 2\frac{g(x_\delta)}{r} \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma) + g(x_\delta) \frac{\tau^2}{n} \text{Tr}(\Sigma^2) + 2\frac{\tau^2}{n} \rho(\Sigma)^2 x_\delta \\ &\leq \left[2\frac{g(x_\delta)}{r} + g(x_\delta) + 2x_\delta \right] \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma) \\ &= t_{\delta,r} \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma), \end{aligned}$$

where $t_{\delta,r}$ is defined as

$$t_{\delta,r} = 2\frac{g(x_\delta)}{r} + g(x_\delta) + 2x_\delta. \quad (29)$$

This entails that $\Psi_{\delta,r} = 0$ and $\hat{\beta}_{\delta,r} = 0$. Using (28) and taking advantage of $\sigma \in \mathcal{S}_L(g(x_\delta)/r)$, we hence get

$$\begin{aligned} \frac{1}{n} \|X\hat{\beta}_\delta - X\beta\|^2 &\leq \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + \frac{2}{n} \|X\bar{\beta}\|^2 \\ &= \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + 2 \left(\frac{\sigma^T \sigma}{\sigma^T \Sigma \sigma} \right)^2 \sigma^T \Sigma \sigma \\ &\leq \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + 2\frac{g(x_\delta)}{r} \frac{\tau^2}{n} \frac{\rho(\Sigma) \text{Tr}(\Sigma)}{\lambda^2}. \end{aligned} \quad (30)$$

This concludes the proof in this first specific regime.

B.2. High signal case (Case ii)

For any $C \in \mathbb{R}^+$, let us define

$$\mathcal{S}_H(C) = \{\sigma \in \mathbb{R}^p, \sigma^T \Sigma \sigma \geq Cp_n\} \quad \text{where} \quad p_n = \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma).$$

We assume in this section that $\sigma \in \mathcal{S}_H(h_{\delta,r})$ where

$$h_{\delta,r} = 2(t_{\delta,r} + g(x_\delta) + 4x_\delta).$$

In this specific regime, $\Psi_{\delta,r} = 1$ on the event \mathcal{A}_δ . Indeed, using first (17) and (21), we obtain

$$\hat{\sigma}^T \Sigma \hat{\sigma} \geq \sigma^T \Sigma \sigma - g(x_\delta) \frac{\tau^2}{n} \text{Tr}(\Sigma^2) - 2 \left(4x_\delta \frac{\tau^2}{n} \rho(\Sigma)^2 \right)^{1/2} (\sigma^T \Sigma \sigma / 2)^{1/2}$$

Since for all $a, b \in \mathbb{R}$, $a^2 - 2ab \geq -b^2$, we get

$$\begin{aligned} \hat{\sigma}^T \Sigma \hat{\sigma} &\geq \frac{1}{2} \sigma^T \Sigma \sigma - g(x_\delta) \frac{\tau^2}{n} \text{Tr}(\Sigma^2) - 4x_\delta \frac{\tau^2}{n} \rho(\Sigma)^2 \\ &\geq \frac{h_{\delta,r}}{2} \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma) - g(x_\delta) \frac{\tau^2}{n} \text{Tr}(\Sigma^2) - 4x_\delta \frac{\tau^2}{n} \rho(\Sigma)^2 \\ &\geq \left(\frac{h_{\delta,r}}{2} - g(x_\delta) - 4x_\delta \right) \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma) \\ &= t_{\delta,r} \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma), \end{aligned}$$

according to the definition of $h_{\delta,r}$. This entails in particular that, on the event \mathcal{A}_δ , $\Psi_{\delta,r} = 1$ and

$$\hat{\beta}_{\delta,r} = \frac{\hat{\sigma}^T \hat{\sigma}}{\hat{\sigma}^T \Sigma \hat{\sigma}} \hat{\sigma} = \hat{\lambda}^{-1} \hat{\sigma} = \hat{\beta}_{PLS}.$$

According to (28), we have

$$\frac{1}{n} \|X(\hat{\beta}_{\delta,r} - \beta)\|^2 \leq \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + \frac{2}{n} \|X(\hat{\beta} - \bar{\beta})\|^2.$$

In the following, we study the second term in the r.h.s. of the previous equality. First remark that

$$\hat{\beta}_{PLS} - \bar{\beta} = \hat{\lambda}^{-1} \hat{\sigma} - \lambda^{-1} \sigma = \hat{\lambda}^{-1} (\hat{\sigma} - \sigma) - \hat{\lambda}^{-1} \frac{\hat{\sigma}^T \Sigma \hat{\sigma} - \sigma^T \Sigma \sigma}{\sigma^T \Sigma \sigma} \sigma + \frac{\hat{\sigma}^T \hat{\sigma} - \sigma^T \sigma}{\sigma^T \Sigma \sigma} \sigma.$$

It yields

$$\frac{2}{n} \|X(\bar{\beta} - \hat{\beta}_{PLS})\|^2 \leq \frac{4}{n} \hat{\lambda}^{-2} \|X(\hat{\sigma} - \sigma)\|^2 + 8 \hat{\lambda}^{-2} \frac{(\hat{\sigma}^T \Sigma \hat{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma} + 8 \frac{(\hat{\sigma}^T \hat{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma}. \quad (31)$$

Remark that $\hat{\lambda}^{-2} = (\hat{\lambda}^{-1} \lambda)^2 \lambda^{-2}$. The term $(\hat{\lambda}^{-1} \lambda)$ illustrates the ratio between the estimation of the norm of the first PLS component and its theoretical value. Using Lemma A.4, still on the event \mathcal{A}_δ , we have

$$\hat{\lambda}^{-1} \lambda \leq C_{\delta,r},$$

where $C_{\delta,r}$ defined in (26). Consequently,

$$\frac{2}{n} \|X(\bar{\beta} - \hat{\beta}_{PLS})\|^2 \leq 4C_{\delta,r}^2 \frac{1}{n} \lambda^{-2} \|X(\hat{\sigma} - \sigma)\|^2 + 8C_{\delta,r}^2 \lambda^{-2} \frac{(\hat{\sigma}^T \Sigma \hat{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma} + 8 \frac{(\hat{\sigma}^T \hat{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma}.$$

A direct application of Lemma A.4 entails that

$$\frac{2}{n} \|X(\bar{\beta} - \hat{\beta})\|^2 \leq 4C_{\delta,r}^2 \lambda^{-2} \mathbf{T}_3(x_\delta) + 8C_{\delta,r}^2 \lambda^{-2} \frac{\mathbf{T}_2(x_\delta)^2}{\sigma^T \Sigma \sigma} + 8 \frac{\mathbf{T}_1(x_\delta)^2}{\sigma^T \Sigma \sigma}. \quad (32)$$

Using (17) and the fact that $\sigma \in \mathcal{S}_H(h_{\delta,r})$, we first get

$$\begin{aligned} \frac{\mathbf{T}_2(x_\delta)^2}{\sigma^T \Sigma \sigma} &\leq 2g(x_\delta)^2 \left(\frac{\tau^2}{n} \text{Tr}(\Sigma^2) \right)^2 \times \frac{1}{\sigma^T \Sigma \sigma} + 16x_\delta \frac{\tau^2}{n} \rho^2(\Sigma), \\ &\leq 2 \frac{g(x_\delta)^2}{h_{\delta,r}} \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma) + 16x_\delta \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma), \\ &= \left(2 \frac{g(x_\delta)^2}{h_{\delta,r}} + 16x_\delta \right) \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma). \end{aligned} \quad (33)$$

Similarly, using (16) and $\sigma \in \mathcal{S}_H(h_{\delta,r})$, we obtain

$$\begin{aligned} \frac{\mathbf{T}_1(x_\delta)^2}{\sigma^T \Sigma \sigma} &\leq 2g(x_\delta)^2 \left(\frac{\tau^2}{n} \text{Tr}(\Sigma) \right)^2 \times \frac{1}{\sigma^T \Sigma \sigma} + 16x_\delta \frac{\tau^2}{n} \rho(\Sigma) \times \lambda^{-1}, \\ &\leq 2 \frac{g(x_\delta)^2}{h_{\delta,r}} \frac{\tau^2}{n} \text{Tr}(\Sigma) \times \frac{1}{\rho(\Sigma)} + 16x_\delta \frac{\tau^2}{n} \rho(\Sigma) \times \lambda^{-1}, \\ &\leq \left(2 \frac{g(x_\delta)^2}{h_{\delta,r}} + 16x_\delta \right) \frac{\tau^2}{n} \text{Tr}(\Sigma) \times \lambda^{-1}, \end{aligned} \quad (34)$$

since $\lambda \leq \rho(\Sigma)$. Gathering (32), (33) and (34), we obtain

$$\begin{aligned} \frac{2}{n} \|X(\bar{\beta} - \hat{\beta})\|^2 &\leq \left(4C_{\delta,r}^2 + 16C_{\delta,r} \frac{g(x_\delta)^2}{h_{\delta,r}} + 128C_{\delta,r}x_\delta \right) \lambda^{-2} \frac{\tau^2}{n} \rho(\Sigma) \text{Tr}(\Sigma) \\ &\quad + \left(16 \frac{g(x_\delta)^2}{h_{\delta,r}} + 128x_\delta \right) \frac{\tau^2}{n} \text{Tr}(\Sigma) \times \lambda^{-1}, \end{aligned}$$

which provides a bound in this regime.

B.3. Intermediate case (Case iii)

We finish the proof with the scenario where $\sigma \in \mathcal{S}_I$ with

$$\mathcal{S}_I = \left\{ \sigma \in \mathbb{R}^p, \frac{g(x_\delta)}{r} p_n \leq \sigma^T \Sigma \sigma \leq h_{\delta,r} p_n \right\}.$$

Remark that \mathcal{S}_I depends on δ and r , but we omit the dependence in the notation. We stress that $\mathcal{S}_I = \mathcal{S}_L(g(x_\delta)/r)^c \cap \mathcal{S}_H(h_\delta)^c$. This section hence covers all the situations that have not been considered before. The fact that $\sigma^T \Sigma \sigma$ is both bounded from above and below allows to control the risk in prediction on \mathcal{A}_δ whatever the value of the indicator function $\Psi_{\delta,r}$ is. Indeed, starting from (28), we have

$$\begin{aligned} \frac{1}{n} \|X(\hat{\beta} - \beta)\|^2 &\leq \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + \frac{2}{n} \|X(\hat{\beta}_{\delta,r} - \bar{\beta})\|^2 \mathbf{1}_{\{\Psi_{\delta,r}=0\}} + \frac{2}{n} \|X(\hat{\beta}_{\delta,r} - \bar{\beta})\|^2 \mathbf{1}_{\{\Psi_{\delta,r} \neq 0\}}, \\ &\leq \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + \frac{2}{n} \|X\bar{\beta}\|^2 + \frac{2}{n} \|X(\hat{\beta}_{PLS} - \bar{\beta})\|^2. \end{aligned}$$

Using $\mathcal{S}_I \subset \mathcal{S}_L(h_{\delta,r})$ and the same algebraic method as in (30), we first obtain

$$\|X\bar{\beta}\|^2 \leq h_{\delta,r} \frac{\tau^2}{n} \frac{\rho(\Sigma) \text{Tr}(\Sigma)}{\lambda^2}.$$

Similarly, using $\mathcal{S}_I \subset \mathcal{S}_H(g(x_\delta)/r)$ and the same bounds as those displayed between (31) and (34), we get

$$\begin{aligned} \frac{2}{n} \|X(\bar{\beta} - \hat{\beta})\|^2 &\leq (4C_{\delta,r}^2 + 16C_{\delta,r} r g(x_\delta) + 128C_{\delta,r} x_\delta) \lambda^{-2} \frac{\tau^2 \rho(\Sigma) \text{Tr}(\Sigma)}{n \lambda^2} \\ &\quad + (16r g(x_\delta) + 128x_\delta) \frac{\tau^2 \text{Tr}(\Sigma)}{n \lambda}. \end{aligned}$$

B.4. End of the proof

Considering the three cases i), ii) and iii), we get

$$\begin{aligned} &\frac{2}{n} \|X(\beta - \hat{\beta}_{\delta,r})\|^2 \\ &\leq \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 \\ &\quad + \left(2 \frac{g(x_\delta)}{r} + 4C_{\delta,r}^2 + 16C_{\delta,r} r g(x_\delta) \max\left(1, \frac{g(x_\delta)}{r h_{\delta,r}}\right) + 128C_{\delta,r} x_\delta \right) \frac{\tau^2 \rho(\Sigma) \text{Tr}(\Sigma)}{n \lambda^2} \\ &\quad + \left(16r g(x_\delta) \max\left(1, \frac{g(x_\delta)}{r h_{\delta,r}}\right) + 128x_\delta \right) \frac{\tau^2 \text{Tr}(\Sigma)}{n \lambda}, \\ &:= \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 + C'_{\delta,r} \frac{\tau^2}{n} \max\left(\frac{\text{Tr}(\Sigma)}{\lambda}, \frac{\rho(\Sigma) \text{Tr}(\Sigma)}{\lambda^2}\right). \end{aligned}$$

As stated previously, the hyperparameter r can be fixed to any value in $(0, 1)$. The calibration of r may be used to improve the constants obtained in the above bound. As we are not looking for optimal constants, we can choose an arbitrary value of r . [Theorem 3.1](#) follows considering e.g. $r = 1/2$, that is, $\hat{\beta}_\delta = \hat{\beta}_{\delta,1/2}$, $C_\delta = C'_{\delta,1/2}$.

C. Technical results - The sparse case

This section contains technical results that will be useful for the proofs of the results presented in [Section 4](#). To this end, we need some additional notation. For any $x \in \mathbb{R}$, we write

$$\text{sgn}(x) = \begin{cases} 1 & \text{if } x > 0, \\ -1 & \text{if } x < 0, \\ 0 & \text{if } x = 0. \end{cases}$$

and $x_+ = x \mathbf{1}_{\{x \geq 0\}}$. In the next two parts, we will write $|\mathcal{J}_0| = \text{Tr}(\Sigma_{\mathcal{J}_0})$ according to [Assumption A.1](#).

C.1. Deviation inequalities

We start with a proposition that lays out the estimator $\hat{\beta}_{sPLS}$ explicitly. [Durif et al. \(2017\)](#) provide a proof with a closed form, for the sake of clarity we reproduce a proof below.

Proposition C.1. *The solution \tilde{w}_1 of the optimisation problem (7) satisfies*

$$\tilde{w}_1 = \frac{\tilde{\sigma}}{\|\tilde{\sigma}\|_2} \quad \text{with} \quad \tilde{\sigma}_j = \text{sgn}(\hat{\sigma}_j)(|\hat{\sigma}_j| - \mu)_+ \quad \forall j \in \{1, \dots, p\}.$$

Moreover, $\hat{\beta}_{sPLS}$ defined in (4) with $W = \tilde{w}_1$ can be written as

$$\hat{\beta}_{sPLS} = \frac{\tilde{\sigma}^T \hat{\sigma}}{\tilde{\sigma}^T \Sigma \tilde{\sigma}} \times \tilde{\sigma}.$$

Proof. Applying the method of Lagrange multipliers, the optimisation problem (7) becomes

$$\underset{w \in \mathbb{R}^p, \nu > 0}{\text{argmin}} \left\{ -\hat{\sigma}^T w + \nu(\|w\|_2^2 - 1) + \mu\|w\|_1 \right\} \quad (35)$$

This problem writes as $\min\{f(w) + g(w)\}$ with $f : w \mapsto -\hat{\sigma}^T w + \nu(\|w\|_2^2 - 1)$ a differentiable convex function and $g : w \mapsto \mu\|w\|_1$ a convex function. Then, w^* is a global optimal solution of (35) if and only if

$$-\nabla f(w^*) \in \partial g(w^*),$$

where ∂g is the sub-differential of the function g (see, e.g., (Bach et al., 2012) for more details). This condition can be written as

$$-\frac{\nabla f}{\mu}(w^*) \in \partial\|\cdot\|_1(w^*) \quad \Leftrightarrow \quad \frac{\hat{\sigma} - \rho w^*}{\mu} \in \partial\|\cdot\|_1(w^*) \quad (36)$$

where in the formula above, $\|\cdot\|_1$ denotes the function $w \mapsto \|w\|_1$. If $w_i^* = 0$, condition (36) leads to $\hat{\sigma}_i \in [-\mu, \mu]$. If $w_i^* > 0$, we obtain $\hat{\sigma}_i - \nu w_i^* = \mu$, which implies $w_i^* = \frac{\hat{\sigma}_i - \mu}{\nu}$. In a similar way for $w_i^* < 0$ we have $w_i^* = \frac{\hat{\sigma}_i + \mu}{\nu}$. We deduce that if $w_i^* \neq 0$, then $\text{sgn}(w_i^*) = \text{sgn}(\hat{\sigma}_i)$.

These computations lead to the closed form

$$w_i^* = \frac{(|\hat{\sigma}_i| - \mu)_+ \text{sgn}(\hat{\sigma}_i)}{\nu} \quad \forall i \in \{1, \dots, p\}.$$

The solution $\tilde{w}_1 = (w_1^*, \dots, w_p^*)$ is then chosen with ν such that the solution has a unitary norm. We derive the formula for $\hat{\beta}_{sPLS}$ thanks to Algorithm 1 with weight \tilde{w}_1 . \square

The following proposition makes precise the relationship between the $\hat{\sigma}_j$ and the threshold μ .

Proposition C.2. *For any $0 < \delta < 1$, let \mathcal{M}_δ the event defined as*

$$\mathcal{M}_\delta = \bigcap_{j=1}^p \left\{ |\hat{\sigma}_j - \sigma_j| \leq \frac{\mu}{2} \right\},$$

with μ as defined in (10). Then, under Assumption A.1,

$$\mathbb{P}(\mathcal{M}_\delta) \geq 1 - \frac{\delta}{2}.$$

Proof. Using Lemma A.1, $\hat{\sigma}_j \sim \mathcal{N}(\sigma_j, \frac{\tau^2}{n})$ for all $j \in \{1, \dots, p\}$ since Σ is normalized according to

Assumption A.1. We get

$$\mathbb{P}(|\hat{\sigma}_j - \sigma_j| > \mu) \leq \frac{\delta}{2p} \quad \forall j \in \{1, \dots, p\}.$$

Then,

$$\mathbb{P}(\mathcal{M}_\delta^c) = \mathbb{P}\left(\bigcup_{j=1}^p \left\{|\hat{\sigma}_j - \sigma_j| > \frac{\mu}{2}\right\}\right) \leq \sum_{j=1}^p \mathbb{P}\left(|\hat{\sigma}_j - \sigma_j| > \frac{\mu}{2}\right) \leq \frac{\delta}{2}.$$

□

In the following, we denote by $\hat{\mathcal{J}}$ the support of $\tilde{\sigma}$, namely

$$\hat{\mathcal{J}} = \{j \in \{1, \dots, p\} \text{ s.t. } |\hat{\sigma}_j| > \mu\}.$$

The exponent C will denote the complementary set *i.e.* for any set \mathcal{I} , $\mathcal{I}^C = \{j \in \{1, \dots, p\}, j \notin \mathcal{I}\}$. On the event \mathcal{M}_δ introduced above, we are able to localize $\hat{\mathcal{J}}$ as shown in the next lemma.

Lemma C.3. *On the event \mathcal{M}_δ we have,*

$$\mathcal{J}_{0,2} \subset \hat{\mathcal{J}} \subset \mathcal{J}_{0,1} \subset \mathcal{J}_0,$$

where

$$\mathcal{J}_{0,1} = \left\{j \in \{1, \dots, p\} \text{ s.t. } |\sigma_j| > \frac{\mu}{2}\right\} \quad \text{and} \quad \mathcal{J}_{0,2} = \{j \in \{1, \dots, p\} \text{ s.t. } |\sigma_j| > 2\mu\}.$$

Proof. On the event \mathcal{M}_δ , for any $j \in \{1, \dots, p\}$,

$$|\sigma_j| > 2\mu \quad \Rightarrow \quad |\hat{\sigma}_j| > \mu.$$

This entails that $\mathcal{J}_{0,2} \subset \hat{\mathcal{J}}$. Similarly, on the event \mathcal{M}_δ , for any $j \in \{1, \dots, p\}$,

$$|\hat{\sigma}_j| > \mu \quad \Rightarrow \quad |\sigma_j| > \frac{\mu}{2},$$

which implies $\hat{\mathcal{J}} \subset \mathcal{J}_{0,1}$. □

By employing a similar line of reasoning as in [Proposition A.3](#), we introduce the following quantities and events:

$$\mathbf{T}_{0,1}(x) = g(x) \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}) + 2\sqrt{2} \sqrt{\frac{\tau^2}{n}} \rho(\Sigma_{\mathcal{J}_0})^{\frac{1}{2}} \sqrt{x} \|\sigma\|_2, \quad (37)$$

$$\mathbf{T}_{0,2}(x) = g(x) \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}), \quad (38)$$

$$\mathbf{T}_{0,3}(x) = g(x) \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}^2), \quad (39)$$

with

$$g(x) = 1 + 2x + 2\sqrt{x}, \quad \forall x \in \mathbb{R}. \quad (40)$$

The following proposition will be the core of the proof for the sparse case. It provides deviation results on the main quantities of interest.

Proposition C.4. *For any $0 < \delta < 1$, let $(\mathcal{B}_{i,\delta})_{i=1}^2$ the events respectively defined as*

$$\mathcal{B}_{1,\delta} = \{|\widehat{\sigma}_{\mathcal{J}_0}^T \widehat{\sigma}_{\mathcal{J}_0} - \sigma^T \sigma| \leq \mathbf{T}_{0,1}(x_{0,\delta})\}, \quad (41)$$

$$\mathcal{B}_{2,\delta} = \{(\widehat{\sigma}_{\mathcal{J}_0} - \sigma)^T (\widehat{\sigma}_{\mathcal{J}_0} - \sigma) \leq \mathbf{T}_{0,2}(x_{0,\delta})\}, \quad (42)$$

$$\text{and } \mathcal{B}_{3,\delta} = \{(\widehat{\sigma}_{\mathcal{J}_0} - \sigma)^T \Sigma (\widehat{\sigma}_{\mathcal{J}_0} - \sigma) \leq \mathbf{T}_{0,3}(x_{0,\delta})\}, \quad (43)$$

with $x_{0,\delta} = \ln(10/\delta)$. Then,

$$\mathbb{P}(\mathcal{B}_\delta) \geq 1 - \frac{\delta}{2} \quad \text{where } \mathcal{B}_\delta := \mathcal{B}_{1,\delta} \cap \mathcal{B}_{2,\delta} \cap \mathcal{B}_{3,\delta}.$$

Proof. The proof is a straightforward generalization of [Proposition A.3](#) and is thus omitted. \square

Using a simple bound, we can easily prove that $\mathbb{P}(\mathcal{M}_\delta \cap \mathcal{B}_\delta) \geq 1 - \delta$.

C.2. Control of $\widetilde{\lambda}$

In the following, we introduce

$$\widetilde{\lambda} = \frac{\widetilde{\sigma}^T \Sigma \widetilde{\sigma}}{\widetilde{\sigma}^T \widehat{\sigma}}.$$

The purpose of this section and of the following proposition is to prove that the ratio $\widetilde{\lambda}^{-1} \lambda$ is controlled with high probability by a constant. The control is acquired under [Assumption A.2](#), that is, when the amount of signal on the component is high enough. It yields, hence, the calibration of the parameter $d_{\delta,p}$ in [Assumption A.2](#).

Proposition C.5. *Assume that [Assumption A.2](#) is satisfied, namely that*

$$\sigma^T \Sigma \sigma > d_{\delta,p} \frac{\tau^2}{n} \rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0}),$$

with $d_{\delta,p} = 4g(x_{0,\delta}) + 192 \ln(\frac{2p}{\delta})$. Then, on the event $\mathcal{M}_\delta \cap \mathcal{B}_\delta$,

$$\widetilde{\lambda}^{-1} \lambda \leq F,$$

with $F = 112$.

Proof. First, remark that

$$\widetilde{\lambda}^{-1} \lambda = \frac{\widetilde{\sigma}^T \widehat{\sigma}}{\sigma^T \sigma} \times \frac{\sigma^T \Sigma \sigma}{\widetilde{\sigma}^T \Sigma \widetilde{\sigma}}.$$

To start the proof, we focus our attention on the first ratio in the r.h.s. of the previous equality. Introduce

$$S_j = \text{sgn}(\widehat{\sigma}_j) \quad \forall j \in \{1, \dots, p\}.$$

We remark that, on the event $\mathcal{M}_\delta \cap \mathcal{B}_\delta$,

$$|\tilde{\sigma}^T \hat{\sigma}| \leq \sum_{j \in \hat{\mathcal{J}}} |\tilde{\sigma}_j \hat{\sigma}_j| \leq \sum_{j \in \hat{\mathcal{J}}} |(\hat{\sigma}_j - \mu S_j) \hat{\sigma}_j| \leq \hat{\sigma}_{\mathcal{J}_0}^T \hat{\sigma}_{\mathcal{J}_0} + \mu \sum_{j \in \hat{\mathcal{J}}} |\hat{\sigma}_j|,$$

since $\hat{\mathcal{J}} \subset \mathcal{J}_{0,1} \subset \mathcal{J}_0$ according to [Lemma C.3](#). Using these inclusions and the Cauchy-Schwarz inequality, we get

$$|\tilde{\sigma}^T \hat{\sigma}| \leq \hat{\sigma}_{\mathcal{J}_0}^T \hat{\sigma}_{\mathcal{J}_0} + \sqrt{\mu^2 |\mathcal{J}_0|} \sqrt{\hat{\sigma}_{\mathcal{J}_0}^T \hat{\sigma}_{\mathcal{J}_0}} \leq 2 \hat{\sigma}_{\mathcal{J}_0}^T \hat{\sigma}_{\mathcal{J}_0} + \mu^2 |\mathcal{J}_0|.$$

[Proposition C.4](#) implies that

$$|\tilde{\sigma}^T \hat{\sigma}| \leq 2\sigma^T \sigma + 2\mathbf{T}_{0,1}(x_{0,\delta}) + \mu^2 |\mathcal{J}_0|.$$

Using the rough bound $\sigma^T \Sigma \sigma \leq \rho(\Sigma_{\mathcal{J}_0}) \sigma^T \sigma$, [Assumption A.2](#) with $d_{\delta,p} = 4g(x_{0,\delta}) + 192 \ln(\frac{2p}{\delta})$ yields

$$\frac{1}{4} \sigma^T \sigma \geq \left(\frac{\tau^2}{n} g(x_{0,\delta}) + 6\mu^2 \right) \text{Tr}(\Sigma_{\mathcal{J}_0}).$$

We obtain

$$\begin{aligned} \tilde{\sigma}^T \hat{\sigma} &\leq 2\sigma^T \sigma + 2g(x_{0,\delta}) \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}) + 4\sqrt{2} \sqrt{\frac{\tau^2}{n} g(x_{0,\delta})} \sqrt{\text{Tr}(\Sigma_{\mathcal{J}_0})} \|\sigma\|_2 + \mu^2 \text{Tr}(\Sigma_{\mathcal{J}_0}) \\ &\leq 2\sigma^T \sigma + \frac{1}{2} \sigma^T \sigma + 2\sqrt{2} \sigma^T \sigma + \frac{1}{24} \sigma^T \sigma \\ &\leq 7\sigma^T \sigma. \end{aligned} \tag{44}$$

Now, we turn our attention to the ratio $\sigma^T \Sigma \sigma / \tilde{\sigma}^T \Sigma \tilde{\sigma}$. We first bound the quantity using the set $\hat{\mathcal{J}}$ as follows:

$$\begin{aligned} \tilde{\sigma}^T \Sigma \tilde{\sigma} &= \tilde{\sigma}_{\hat{\mathcal{J}}}^T \Sigma \tilde{\sigma}_{\hat{\mathcal{J}}} \\ &= (\hat{\sigma} - \mu S)_{\hat{\mathcal{J}}}^T \Sigma (\hat{\sigma} - \mu S)_{\hat{\mathcal{J}}} \\ &= \hat{\sigma}_{\hat{\mathcal{J}}}^T \Sigma \hat{\sigma}_{\hat{\mathcal{J}}} + \mu^2 S_{\hat{\mathcal{J}}}^T \Sigma S_{\hat{\mathcal{J}}} - 2\mu \hat{\sigma}_{\hat{\mathcal{J}}}^T \Sigma S_{\hat{\mathcal{J}}} \\ &\geq \frac{1}{2} \hat{\sigma}_{\hat{\mathcal{J}}}^T \Sigma \hat{\sigma}_{\hat{\mathcal{J}}} - \mu^2 S_{\hat{\mathcal{J}}}^T \Sigma S_{\hat{\mathcal{J}}} \\ &\geq \frac{1}{2} \hat{\sigma}_{\hat{\mathcal{J}}}^T \Sigma \hat{\sigma}_{\hat{\mathcal{J}}} - \mu^2 \rho(\Sigma_{\hat{\mathcal{J}}}) |\hat{\mathcal{J}}| \\ &\geq \frac{1}{2} \hat{\sigma}_{\hat{\mathcal{J}}}^T \Sigma \hat{\sigma}_{\hat{\mathcal{J}}} - \mu^2 \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|, \end{aligned}$$

where the last inequality comes from the inclusion $\hat{\mathcal{J}} \subset \mathcal{J}_0$. Now we want to find a bound relying

only on the set \mathcal{J}_0 . Note that

$$\begin{aligned}
\hat{\sigma}_{\hat{\mathcal{J}}}^T \Sigma \hat{\sigma}_{\hat{\mathcal{J}}} &= (\hat{\sigma}_{\mathcal{J}_0} - \hat{\sigma}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c})^T \Sigma (\hat{\sigma}_{\mathcal{J}_0} - \hat{\sigma}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c}) \\
&= \hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0} - 2 \hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c} + \hat{\sigma}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c}^T \Sigma \hat{\sigma}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c} \\
&\geq \frac{\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0}}{2} - \hat{\sigma}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c}^T \Sigma \hat{\sigma}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c} \\
&\geq \frac{\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0}}{2} - \mu^2 \mathbf{1}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c}^T \Sigma \mathbf{1}_{\mathcal{J}_0 \cap \hat{\mathcal{J}}^c} \\
&\geq \frac{\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0}}{2} - \mu^2 \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|.
\end{aligned}$$

We deduce the following inequality:

$$\begin{aligned}
\tilde{\sigma}^T \Sigma \tilde{\sigma} &\geq \frac{\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0}}{4} - \frac{\mu^2}{2} \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0| - \mu^2 \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0| \\
&\geq \frac{\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0}}{4} - \frac{3\mu^2}{2} \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0| \\
&\geq \frac{1}{4} (\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0} - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0|).
\end{aligned} \tag{45}$$

Finally, we obtain a bound that does not depend on \mathcal{J}_0 . Indeed,

$$\begin{aligned}
&\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0} - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0| \\
&= (\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + 2\sigma^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + \sigma^T \Sigma \sigma - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0| \\
&\geq (\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_0} - \frac{1}{2} \sigma^T \Sigma \sigma - 2(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + \sigma^T \Sigma \sigma - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0| \\
&\geq \frac{1}{2} \sigma^T \Sigma \sigma - (\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_0} - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0|.
\end{aligned}$$

Using the fact that we are on the event \mathcal{B}_δ , we have

$$\begin{aligned}
\hat{\sigma}_{\mathcal{J}_0}^T \Sigma \hat{\sigma}_{\mathcal{J}_0} - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0| &\geq \frac{1}{2} \sigma^T \Sigma \sigma - \mathbf{T}_{0,2}(x_{0,\delta}) - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0| \\
&\geq \frac{1}{2} \sigma^T \Sigma \sigma - g(x_{0,\delta}) \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}^2) - 6\rho(\Sigma_{\mathcal{J}_0}) \mu^2 |\mathcal{J}_0|, \\
&\geq \frac{1}{2} \sigma^T \Sigma \sigma - \left(\frac{\tau^2}{n} g(x_{0,\delta}) + 6\mu^2 \right) \rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0}), \\
&\geq \frac{1}{4} \sigma^T \Sigma \sigma,
\end{aligned} \tag{46}$$

where we have used [Assumption A.2](#) for the last inequality, with $d_{\delta,p} = 4g(x_{0,\delta}) + 192 \ln(\frac{2p}{\delta})$.

We deduce from (45) and (46) that

$$\tilde{\sigma}^T \Sigma \tilde{\sigma} \geq \frac{1}{16} \sigma^T \Sigma \sigma. \tag{47}$$

Inequalities (44) and (47) lead to

$$\tilde{\lambda}^{-1} \lambda \leq 7 \times 16,$$

which concludes the proof. \square

D. Proof of the results displayed in Section 4

D.1. Proof of Theorem 4.1

From now on we work on the event $\mathcal{M}_\delta \cap \mathcal{B}_\delta$. First

$$\frac{1}{n} \|X \widehat{\beta}_{sPLS} - X\beta\|^2 \leq \frac{2}{n} \|X(\widehat{\beta}_{sPLS} - \bar{\beta})\|^2 + \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2,$$

with $\bar{\beta} = \frac{\sigma^T \sigma}{\sigma^T \Sigma \sigma} \sigma$. We use again the segmentation proposed in the proof of Theorem 3.1,

$$\widehat{\beta}_{sPLS} - \bar{\beta} = \tilde{\lambda}^{-1} \tilde{\sigma} - \lambda^{-1} \sigma = \tilde{\lambda}^{-1} (\tilde{\sigma} - \sigma) - \tilde{\lambda}^{-1} \frac{\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma}{\sigma^T \Sigma \sigma} \sigma + \frac{\tilde{\sigma}^T \widehat{\sigma} - \sigma^T \sigma}{\sigma^T \Sigma \sigma} \sigma.$$

It yields

$$\frac{2}{n} \|X(\bar{\beta} - \widehat{\beta}_{sPLS})\|^2 \leq \frac{4}{n} \tilde{\lambda}^{-2} \|X(\tilde{\sigma} - \sigma)\|^2 + 8 \tilde{\lambda}^{-2} \frac{(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma} + 8 \frac{(\tilde{\sigma}^T \widehat{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma}. \quad (48)$$

Remark that $\tilde{\lambda}^{-2} = (\tilde{\lambda}^{-1} \lambda)^2 \lambda^{-2}$. The term $(\tilde{\lambda}^{-1} \lambda)$ represents the ratio between the estimation of the norm of the first PLS component and its true value. Using Proposition C.5,

$$\tilde{\lambda}^{-1} \lambda \leq 112 := F.$$

Consequently,

$$\frac{2}{n} \|X(\bar{\beta} - \widehat{\beta}_{sPLS})\|^2 \leq \underbrace{4F^2 \frac{1}{n} \lambda^{-2} \|X(\tilde{\sigma} - \sigma)\|^2}_I + \underbrace{8F^2 \lambda^{-2} \frac{(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma}}_{II} + \underbrace{8 \frac{(\tilde{\sigma}^T \widehat{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma}}_{III}. \quad (49)$$

The proof naturally falls into three parts.

D.1.1. Study of term I

First, we focus on the deviation of

$$\frac{1}{n} \|X(\tilde{\sigma} - \sigma)\|^2 = (\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma).$$

Let us decompose this quantity as follows,

$$(\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma) \leq 2(\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}}^T \Sigma (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}} + 2(\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}^c}^T \Sigma (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}^c}, \quad (50)$$

with indexes set $\mathcal{J}_{0,1}$ defined in Lemma C.3. We consider separately the two terms on the right hand side.

Using the inclusion $\mathcal{J}_{0,1} \subset \mathcal{J}_0$,

$$\begin{aligned}
& (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}}^T \Sigma (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}} \\
& \leq 2(\hat{\sigma} - \sigma)_{\mathcal{J}_{0,1}}^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_{0,1}} + 2\mu^2 S_{\mathcal{J}_{0,1}}^T \Sigma S_{\mathcal{J}_{0,1}} \\
& = 2(\hat{\sigma}_{\mathcal{J}_0} - \sigma_{\mathcal{J}_0} - (\hat{\sigma} - \sigma)_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c})^T \Sigma (\hat{\sigma}_{\mathcal{J}_0} - \sigma_{\mathcal{J}_0} - (\hat{\sigma} - \sigma)_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c}) + 2\mu^2 S_{\mathcal{J}_{0,1}}^T \Sigma S_{\mathcal{J}_{0,1}} \\
& \leq 4(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T \Sigma_{\mathcal{J}_0} (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + 4(\hat{\sigma} - \sigma)_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c}^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c} + 2\mu^2 \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|.
\end{aligned}$$

As stated in Lemma C.3, on \mathcal{M}_δ , $\mathcal{J}_{0,1}^C \subset \hat{\mathcal{J}}^C$. Hence, writing $\mathbf{1} = (1, \dots, 1)^T \in \mathbb{R}^p$,

$$(\hat{\sigma} - \sigma)_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c}^T \Sigma (\hat{\sigma} - \sigma)_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c} \leq (\mu \mathbf{1} + \frac{\mu}{2} \mathbf{1})_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c}^T \Sigma_{\mathcal{J}_0} (\mathbf{1} \mu + \mathbf{1} \frac{\mu}{2})_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c} \leq \frac{9\mu^2}{4} \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|,$$

since for all $j \in \mathcal{J}_{0,1}^C$, $|\sigma_j| < \mu/2$. Consequently,

$$(\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}}^T \Sigma (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}} \leq 4(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T \Sigma_{\mathcal{J}_0} (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + 11\mu^2 \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|. \quad (51)$$

Similarly,

$$\begin{aligned}
(\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}^C}^T \Sigma (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}^C} & \leq \sigma_{\mathcal{J}_{0,1}^C}^T \Sigma \sigma_{\mathcal{J}_{0,1}^C} \\
& = \sigma_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^C}^T \Sigma_{\mathcal{J}_0} \sigma_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^C} \\
& \leq 2\frac{\mu^2}{4} \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|.
\end{aligned} \quad (52)$$

With inequalities (50), (51) and (52), we obtain

$$(\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma) \leq 8(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T \Sigma_{\mathcal{J}_0} (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + 23\mu^2 \rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|.$$

We bound the first term using inequality (43), since we are working on the event \mathcal{B}_δ . For the second term, we replace μ by its expression. We obtain

$$\frac{(\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma)}{\lambda^2} \leq 8g(x_{0,\delta}) \frac{\tau^2 \text{Tr}(\Sigma_{\mathcal{J}_0}^2)}{n \lambda^2} + 8 \cdot 23 \frac{\tau^2}{n} \ln \left(\frac{2p}{\delta} \right) \frac{\rho(\Sigma_{\mathcal{J}_0}) |\mathcal{J}_0|}{\lambda^2}. \quad (53)$$

Hence, for $0 < \delta < 1/2$,

$$\frac{(\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma)}{\lambda^2} \leq C_I \ln \left(\frac{p}{\delta} \right) \frac{\tau^2 \rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0})}{n \lambda^2}, \quad (54)$$

with C_I a positive constant.

D.1.2. Study of term II

The task is now to bound the term

$$\lambda^{-2} \frac{(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma}.$$

We use the equality

$$\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma = (\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma) + 2\sigma^T \Sigma (\tilde{\sigma} - \sigma).$$

Using the Cauchy Schwarz inequality on the last term,

$$\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma \leq (\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma) + 2\sqrt{\sigma^T \Sigma \sigma} \sqrt{(\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma)}.$$

Then,

$$(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2 \leq 2((\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma))^2 + 8\sigma^T \Sigma \sigma \times (\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma).$$

Using inequality (53) on the first term, we get

$$\frac{((\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma))^2}{\sigma^T \Sigma \sigma \lambda^2} \leq \left(2 \cdot 64 \cdot g(x_{0,\delta})^2 + 2 \cdot 224^2\right) \left(\frac{\tau^2}{n}\right)^2 \frac{\rho(\Sigma_{\mathcal{J}_0})^2 |\mathcal{J}_0|^2}{\sigma^T \Sigma \sigma \lambda^2} \ln \left(\frac{2p}{\delta}\right)^2.$$

Applying inequality (54) on the second term, we get

$$\frac{\sigma^T \Sigma \sigma (\tilde{\sigma} - \sigma)^T \Sigma (\tilde{\sigma} - \sigma)}{\sigma^T \Sigma \sigma \lambda^2} \leq C_I \ln \left(\frac{p}{\delta}\right) \frac{\tau^2}{n} \frac{\rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0})}{\lambda^2}.$$

We bound $1/(\sigma^T \Sigma \sigma)$ in the first term using Assumption A.2. We obtain

$$\begin{aligned} \frac{(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma \lambda^2} &\leq 2(128 g(x_{0,\delta})^2 + 2 \cdot 224^2) d_{\delta,p}^{-1} \ln \left(\frac{2p}{\delta}\right)^2 \frac{\tau^2}{n} \frac{\rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0})}{\lambda^2} \\ &\quad + 8C_I \ln \left(\frac{p}{\delta}\right) \frac{\tau^2}{n} \frac{\rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0})}{\lambda^2}. \end{aligned}$$

When $d_{\delta,p} = 4g(x_{0,\delta}) + 192 \ln(\frac{2p}{\delta})$, we deduce

$$\lambda^{-2} \frac{(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma} \leq C_{II,\delta} \ln \left(\frac{p}{\delta}\right) \frac{\tau^2}{n} \frac{\rho(\Sigma_{\mathcal{J}_0}) \text{Tr}(\Sigma_{\mathcal{J}_0})}{\lambda^2}, \quad (55)$$

where C_{II} is a positive constant.

D.1.3. Study of term III

Finally we focus on the last term,

$$\frac{(\tilde{\sigma}^T \hat{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma}.$$

Using the inclusion $\hat{\mathcal{J}} \subset \mathcal{J}_0$ given by Lemma C.3, we get

$$\begin{aligned} \tilde{\sigma}^T \hat{\sigma} - \sigma^T \sigma &= \tilde{\sigma}_{\mathcal{J}_0}^T \hat{\sigma}_{\mathcal{J}_0} - \sigma_{\mathcal{J}_0}^T \sigma_{\mathcal{J}_0} \\ &= (\tilde{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + \sigma_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + \sigma_{\mathcal{J}_0}^T (\tilde{\sigma} - \sigma)_{\mathcal{J}_0} \\ &\leq \sqrt{(\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma)} \sqrt{(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0}} \\ &\quad + \sqrt{\sigma^T \sigma} \left(\sqrt{(\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma)} + \sqrt{(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0}} \right). \end{aligned}$$

Using the inequality $2ab \leq a^2 + b^2$ for all $a, b \in \mathbb{R}$, it follows that

$$\begin{aligned}
& (\tilde{\sigma}^T \hat{\sigma} - \sigma^T \sigma)^2 \\
& \leq 2(\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma) \times (\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + 4\sigma^T \sigma \times ((\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma) + (\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0}) \\
& \leq ((\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma))^2 + ((\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0})^2 + 4\sigma^T \sigma ((\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma) + (\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0}).
\end{aligned} \tag{56}$$

First, we concentrate our attention on the quantity $(\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma)$. Using [Lemma C.3](#), we obtain

$$\begin{aligned}
(\tilde{\sigma} - \sigma)^T (\tilde{\sigma} - \sigma) & \leq 2(\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}}^T (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}} + 2(\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}^c}^T (\tilde{\sigma} - \sigma)_{\mathcal{J}_{0,1}^c} \\
& \leq 2(\hat{\sigma} - \sigma - \mu S)_{\mathcal{J}_{0,1}}^T (\hat{\sigma} - \sigma - \mu S)_{\mathcal{J}_{0,1}} + 2\sigma_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c}^T \sigma_{\mathcal{J}_0 \cap \mathcal{J}_{0,1}^c} \\
& \leq 4(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + 4\mu^2 S_{\mathcal{J}_0}^T S_{\mathcal{J}_0} + \frac{1}{2}\mu^2 |\mathcal{J}_0| \\
& \leq 4(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + \frac{9}{2}\mu^2 |\mathcal{J}_0|.
\end{aligned}$$

Starting from [\(56\)](#), we obtain

$$\begin{aligned}
(\tilde{\sigma}^T \hat{\sigma} - \sigma^T \sigma)^2 & \leq \left(4(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + \frac{9}{2}\mu^2 |\mathcal{J}_0| \right)^2 + ((\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0})^2 \\
& \quad + 4\sigma^T \sigma \left(5(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} + \frac{9}{2}\mu^2 |\mathcal{J}_0| \right).
\end{aligned}$$

Note that, on the event \mathcal{B}_δ , equation [\(42\)](#) in [Proposition C.4](#) reads as

$$(\hat{\sigma} - \sigma)_{\mathcal{J}_0}^T (\hat{\sigma} - \sigma)_{\mathcal{J}_0} \leq g(x_{0,\delta}) \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}).$$

Hence,

$$\begin{aligned}
(\tilde{\sigma}^T \hat{\sigma} - \sigma^T \sigma)^2 & \leq 33g(x_{0,\delta})^2 \left(\frac{\tau^2}{n} \right)^2 \text{Tr}(\Sigma_{\mathcal{J}_0})^2 + 2 \cdot 36^2 \left(\frac{\tau^2}{n} \right)^2 \ln \left(\frac{2p}{\delta} \right)^2 |\mathcal{J}_0|^2 \\
& \quad + 4\sigma^T \sigma \left(5g(x_{0,\delta}) \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}) + 36 \frac{\tau^2}{n} \ln \left(\frac{2p}{\delta} \right) |\mathcal{J}_0| \right).
\end{aligned}$$

Observe that [Assumption A.2](#) yields

$$\sigma^T \Sigma \sigma = \lambda \sigma^T \sigma \geq \lambda \frac{\tau^2}{n} \text{Tr}(\Sigma_{\mathcal{J}_0}).$$

Hence,

$$\begin{aligned}
\frac{(\tilde{\sigma}^T \hat{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma} & \leq 33 \frac{g(x_{0,\delta})^2 \tau^2}{d_{\delta,p}} \frac{\text{Tr}(\Sigma_{\mathcal{J}_0})}{n \lambda} + 2 \cdot 36^2 \frac{\ln \left(\frac{2p}{\delta} \right)^2 \tau^2}{d_{\delta,p}} \frac{\text{Tr}(\Sigma_{\mathcal{J}_0})}{n \lambda} \\
& \quad + 20g(x_{0,\delta}) \frac{\tau^2}{n} \frac{\text{Tr}(\Sigma_{\mathcal{J}_0})}{\lambda} + 144 \frac{\tau^2}{n} \ln \left(\frac{2p}{\delta} \right) \frac{\text{Tr}(\Sigma_{\mathcal{J}_0})}{\lambda} \\
& \leq C_{III,\delta} \ln \left(\frac{p}{\delta} \right) \frac{\tau^2}{n} \frac{\text{Tr}(\Sigma_{\mathcal{J}_0})}{\lambda}
\end{aligned} \tag{57}$$

where $C_{III,\delta}$ is a positive constant depending only on δ , when $d_{\delta,p} = 4g(x_{0,\delta}) + 192 \ln(\frac{2p}{\delta})$.

D.1.4. End of the proof

Theorem 4.1 follows directly from (49), (54), (55) and (57). It has been proven under **Assumption A.2** when $d_{\delta,p} = 4g(x_{0,\delta}) + 192 \ln(\frac{2p}{\delta})$. Hence, the result still holds considering $d_{\delta,p} = C_0 (\ln(\frac{10}{\delta}) + \ln(\frac{p}{\delta}))$ with $C_0 = 384$.

D.2. Proof of Corollary 4.2

Since the support of σ is \mathcal{J}_0 , we have immediately that

$$\|\sigma_{\mathcal{J}_0^c}\|_1 = 0 < \|\sigma_{\mathcal{J}_0}\|_1.$$

Hence, using **Assumption A.3**, we get that

$$\lambda := \frac{\sigma^T \Sigma \sigma}{\sigma^T \sigma} = \frac{1}{n} \frac{\|X\sigma\|^2}{\|\sigma\|^2} > \frac{1}{\phi}.$$

A direct application of **Theorem 4.1** then indicates that

$$\begin{aligned} \frac{1}{n} \|X(\widehat{\beta}_{sPLS} - \beta)\|^2 &\leq \frac{2}{n} \inf_{v \in [\sigma]} \|X(\beta - v)\|^2 + D_\delta \frac{\tau^2 s}{n} \max\left(\frac{\rho(\Sigma_{\mathcal{J}_0})}{\lambda^2}, \frac{1}{\lambda}\right) \ln\left(\frac{p}{\delta}\right) \\ &\leq \frac{2}{n} \inf_{v \in [\sigma]} \|X(\beta - v)\|^2 + D_\delta \frac{\tau^2 s}{n} \max(\phi^2 \rho(\Sigma_{\mathcal{J}_0}), \phi) \ln\left(\frac{p}{\delta}\right), \end{aligned}$$

which exactly corresponds to the desired results.

D.3. Proof of Theorem 4.3

We first remark that the estimator (13) can be rewritten as

$$\tilde{\beta} = \frac{\tilde{\sigma}^T \tilde{\sigma}}{\tilde{\sigma}^T \Sigma \tilde{\sigma}} \tilde{\sigma} := (\tilde{\lambda}^*)^{-1} \tilde{\sigma} \quad \text{where} \quad \tilde{\lambda}^* = \frac{\tilde{\sigma}^T \Sigma \tilde{\sigma}}{\tilde{\sigma}^T \tilde{\sigma}}.$$

We use similar steps as those used in **Theorem 4.1**. First

$$\frac{1}{n} \|X\tilde{\beta} - X\beta\|^2 \leq \frac{2}{n} \|X(\tilde{\beta} - \bar{\beta})\|^2 + \frac{2}{n} \|X(\bar{\beta} - \beta)\|^2 \quad \text{with} \quad \bar{\beta} = \frac{\sigma^T \sigma}{\sigma^T \Sigma \sigma} \sigma.$$

Then, using the previous segmentation studied in the proofs of **Theorem 3.1** and **Theorem 4.1**, we get

$$\tilde{\beta} - \bar{\beta} = (\tilde{\lambda}^*)^{-1} \tilde{\sigma} - \lambda^{-1} \sigma = (\tilde{\lambda}^*)^{-1} (\tilde{\sigma} - \sigma) - (\tilde{\lambda}^*)^{-1} \frac{\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma}{\tilde{\sigma}^T \Sigma \sigma} \sigma + \frac{\tilde{\sigma}^T \tilde{\sigma} - \sigma^T \sigma}{\sigma^T \Sigma \sigma} \sigma.$$

It yields

$$\frac{2}{n} \|X(\bar{\beta} - \tilde{\beta})\|^2 \leq \frac{4}{n} (\tilde{\lambda}^*)^{-2} \|X(\tilde{\sigma} - \sigma)\|^2 + 8 (\tilde{\lambda}^*)^{-2} \frac{(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma} + 8 \frac{(\tilde{\sigma}^T \tilde{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma}. \quad (58)$$

From now on, we work on the event $\mathcal{M}_\delta \cap \mathcal{B}_\delta$. We first concentrate our attention on $\tilde{\lambda}^*$. Using [Lemma C.3](#), $\hat{\mathcal{J}} \subset \mathcal{J}_0$. Hence $\|\tilde{\sigma}_{\mathcal{J}_0^c}\|_1 = 0 \leq 3\|\tilde{\sigma}_{\mathcal{J}_0}\|_1$, which entails, using [Assumption A.3](#) that

$$\tilde{\lambda}^* := \frac{\tilde{\sigma}^T \Sigma \tilde{\sigma}}{\tilde{\sigma}^T \tilde{\sigma}} = \frac{1}{n} \frac{\|X \tilde{\sigma}\|^2}{\|\tilde{\sigma}\|_2^2} \geq \frac{1}{\phi}.$$

Then inequality (58) immediately leads to

$$\frac{2}{n} \|X(\bar{\beta} - \tilde{\beta})\|^2 \leq \frac{4}{n} \phi^2 \|X(\tilde{\sigma} - \sigma)\|^2 + 8 \phi^2 \frac{(\tilde{\sigma}^T \Sigma \tilde{\sigma} - \sigma^T \Sigma \sigma)^2}{\sigma^T \Sigma \sigma} + 8 \frac{(\tilde{\sigma}^T \tilde{\sigma} - \sigma^T \sigma)^2}{\sigma^T \Sigma \sigma}. \quad (59)$$

A direct application of [Proposition C.4](#) then allows to control each of the terms involved in (59) and then to conclude the proof.

References

- Abdel-Rahman, E. M., O. Mutanga, J. Odindi, E. Adam, A. Odindo, and R. Ismail (2014). “A comparison of partial least squares (PLS) and sparse PLS regressions for predicting yield of Swiss chard grown under different irrigation water sources using hyperspectral data”. In: *Computers and Electronics in Agriculture* 106, pp. 11–19.
- Alsouki, L., L. Duval, C. Marteau, R. E. Haddad, and F. Wahl (2023). “Dual-sPLS: a family of Dual Sparse Partial Least Squares regressions for feature selection and prediction with tunable sparsity; evaluation on simulated and near-infrared (NIR) spectra”. In: *Chemometrics and Intelligent Laboratory system* 237.
- Bach, F., R. Jenatton, J. Mairal, and G. Obozinski (2012). “Optimization with Sparsity-Inducing Penalties”. In: *Foundations and Trends® in Machine Learning* 4.1, pp. 1–106.
- Bickel, P. J., Y. Ritov, and A. B. Tsybakov (2009). “Simultaneous analysis of lasso and Dantzig selector”. In: *Ann. Statist.* 37.4, pp. 1705–1732.
- Cao, K.-A. L., D. Rossouw, C. Robert-Granié, and P. Besse (2008). “A Sparse PLS for Variable Selection when Integrating Omics Data”. In: *Statistical Applications in Genetics and Molecular Biology* 7.1.
- Chun, H. and S. Keleş (2010). “Sparse partial least squares regression for simultaneous dimension reduction and variable selection”. In: *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* 72.1, pp. 3–25.
- Cook, R. D. and L. Forzani (Apr. 2017). “Big data and partial least-squares prediction”. In: *Canadian Journal of Statistics* 46.
- (2019). “Partial least squares prediction in high-dimensional regression”. In: *The Annals of Statistics* 47.2, pp. 884–908.
- Cook, R. D., I. S. Helland, and Z. Su (2013). “Envelopes and partial least squares regression”. In: *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* 75.5, pp. 851–877.

- Cook, R. D., B. Li, and F. Chiaromonte (2010). “Envelope models for parsimonious and efficient multivariate linear regression”. In: *Statistica Sinica*, pp. 927–960.
- Dalalyan, A. S., M. Hebiri, and J. Lederer (2017). “On the prediction performance of the Lasso”. In: *Bernoulli* 23.1, pp. 552–581.
- de Jong, S. (1993). “SIMPLS: An alternative approach to partial least squares regression”. In: *Chemometrics and Intelligent Laboratory Systems* 18.3, pp. 251–263.
- Delaigle, A. and P. Hall (2012). “Methodology and theory for partial least squares applied to functional data”. In:
- Donoho, D. L. and I. M. Johnstone (1994). “Ideal spatial adaptation by wavelet shrinkage”. In: *biometrika* 81.3, pp. 425–455.
- Durif, G., L. Modolo, J. Michaelsson, J. E. Mold, S. Lambert-Lacroix, and F. Picard (2017). “High dimensional classification with combined adaptive sparse PLS and logistic regression”. In: *Bioinformatics* 34.3, pp. 485–493.
- Frank, L. E. and J. Friedman (1993). “A statistical view of some chemometrics regression tools”. In: *Technometrics* 35.2, pp. 109–135.
- Fuentes, J., P. Poncela, and J. Rodríguez (2015). “Sparse partial least squares in time series for macroeconomic forecasting”. In: *Journal of Applied Econometrics* 30.4, pp. 576–595.
- Garthwaite, P. H. (1994). “An interpretation of partial least squares”. In: *Journal of the American Statistical Association* 89.425, pp. 122–127.
- Giraud, C. (2021). *Introduction to high-dimensional statistics*. CRC Press.
- Helland, I. S. (1990). “Partial least squares regression and statistical models”. In: *Scandinavian journal of statistics*, pp. 97–114.
- Krämer, N. and M. Sugiyama (2011). “The degrees of freedom of partial least squares regression”. In: *Journal of the American Statistical Association* 106.494, pp. 697–705.
- Laurent, B., J.-M. Loubes, and C. Marteau (2012). “Non asymptotic minimax rates of testing in signal detection with heterogeneous variances”. In: *Electronic Journal of Statistics* 6, pp. 91–122.
- Lee, D., W. Lee, Y. Lee, and Y. Pawitan (2011). “Sparse partial least-squares regression and its applications to high-throughput data analysis”. In: *Chemometrics and Intelligent Laboratory Systems* 109.1, pp. 1–8.
- Martens, H. and T. Naes (1992). *Multivariate calibration*. John Wiley & Sons.
- Mateos-Aparicio, G. (2011). “Partial Least Squares (PLS) Methods: Origins, Evolution, and Application to Social Sciences”. In: *Communications in Statistics - Theory and Methods* 40.13, pp. 2305–2317.
- Meinshausen, N. (2013). “Sign-constrained least squares estimation for high-dimensional regression”. In: *Electronic Journal of Statistics* 7.none, pp. 1607–1631.
- Naik, P. and C.-L. Tsai (2000). “Partial least squares estimator for single-index models”. In: *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* 62.4, pp. 763–771.
- Palermo, G., P. Piraino, and H.-D. Zucht (2009). “Performance of PLS regression coefficients in selecting variables for each response of a multivariate PLS for omics-type data”. In: *Advances and Applications in Bioinformatics and Chemistry*, pp. 57–70.
- Qin, S. J., Y. Liu, and S. Tang (2023). “Partial least squares, steepest descent, and conjugate gradient for regularized predictive modeling”. In: *AIChE Journal* 69.4, e17992.
- Sawatsky, M. L., M. Clyde, and F. Meek (2015). “Partial least squares regression in the social sciences”. In: *The Quantitative Methods for Psychology* 11.2, pp. 52–62.

- Tibshirani, R. (1996). "Regression Shrinkage and Selection Via the Lasso". In: *Journal of the Royal Statistical Society: Series B (Methodological)* 58.1, pp. 267–288.
- Wold, H. (1975). "Path Models with Latent Variables: The NIPALS Approach". In: *Quantitative Sociology*. Ed. by H. Blalock, A. Aganbegian, F. Borodkin, R. Boudon, and V. Capecchi. International Perspectives on Mathematical and Statistical Modeling. Academic Press, pp. 307–357.
- Wold, S. (1995). "Chemometrics; what do we mean with it, and what do we want from it?" In: *Chemometrics and Intelligent Laboratory Systems* 30.1. In CINC '94 Selected papers from the First International Chemometrics Internet Conference, pp. 109–115.
- Wold, S., M. Sjöström, and L. Eriksson (2001). "PLS-regression: a basic tool of chemometrics". In: *Chemometrics and Intelligent Laboratory Systems* 58.2. PLS Methods, pp. 109–130.
- Yang, T. C., L. S. Aucott, G. G. Duthie, and H. M. Macdonald (2017). "An application of partial least squares for identifying dietary patterns in bone health". In: *Archives of osteoporosis* 12, pp. 1–8.