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# The Solution Path of SLOPE

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## Abstract

The SLOPE estimator has the particularity of having null components (sparsity) and components that are equal in absolute value (clustering). The number of clusters depends on the regularization parameter of the estimator. This parameter can be chosen as a trade-off between interpretability (with a small number of clusters) and accuracy (with a small mean squared error or a small prediction error). Finding such a compromise requires to compute the solution path, that is the function mapping the regularization parameter to the estimator. We provide in this article an algorithm to compute the solution path of SLOPE.

## 1 Introduction

The SLOPE estimator (Sorted L One Penalized Estimator [1, 21]) is defined as a solution to the following convex program:

$$\min_{b \in \mathbb{R}^p} \frac{1}{2} \|y - Xb\|_2^2 + \gamma \sum_{i=1}^p \lambda_i |b|_{\downarrow i}. \quad (1)$$

In (1),  $\lambda_1 > 0$ ,  $\lambda_1 \geq \dots \geq \lambda_p \geq 0$  is a given sequence of penalty parameters,  $\gamma > 0$  is the regularization parameter and  $|b|_{\downarrow 1} \geq \dots \geq |b|_{\downarrow p} \geq 0$  are sorted components of  $b$  in absolute value. The SLOPE estimator generalizes both the LASSO estimator (Least Absolute Shrinkage and Selection Operator [20]) for which  $\lambda_1 = \dots = \lambda_p = 1$ , and the OSCAR estimator (Octagonal Shrinkage and Clustering Algorithm for Regression [3]) for which the sequence  $\lambda_1, \dots, \lambda_p$  is arithmetic.

The SLOPE estimator is gaining popularity among statisticians due to its relevant properties such as false discovery rate control [1] and dimension reduction of the regression model. The latter property comes from the structure of the solutions to the optimization problem (1), which have null components (sparsity) as well as components equal in absolute value (clustering) [14, 6, 3]<sup>1</sup>. When  $y$  represents the random response of a linear regression model, sparsity has a well-known statistical interpretation: identification of relevant explanatory variables. Clustering also has a statistical interpretation when the design matrix  $X$  is standardized: the explanatory variables having the same regression coefficient have the same impact on the response [15]. On the other hand, without restriction on the design matrix, for a categorical variable having different levels, the equal regression coefficients represent levels that can be grouped together [17, 10]. Therefore, SLOPE estimator can identify relevant explanatory variables, group explanatory variables having the same impact on the response and, more generally, reduce the dimension of the regression model.

The solution path gives the solution of a penalized optimization problem with respect to the regularization parameter  $\gamma > 0$ . For the LASSO, this path shows that the number of explanatory

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<sup>1</sup>When  $X$  is an orthogonal matrix, the solution to the problem (1) is explicit and its sparsity and clustering properties are straightforward [1, 5, 16, 19]).

variables selected by this estimator tends to decrease when the regularization parameter becomes large (see e.g. [9, 13]). Adjusting the regularization parameter  $\gamma$  allows a compromise between selecting a small number of explanatory variables and constructing an accurate estimator. Similarly, the construction of the solution path of SLOPE is useful to adjust the regularization parameter to have a good trade-off between interpretability (by selecting a small number of clusters of explanatory variables) and accuracy (with a small mean squared error or a small prediction error).

In this article, given a sequence of penalty parameters  $\lambda_1 > \dots > \lambda_p > 0$ , we prove that the solution path of SLOPE is piecewise linear on  $(0, +\infty)$ , we characterize its affine components, and we provide an algorithm to compute the path.<sup>2</sup>

## 2 Basic notions on SLOPE

### 2.1 Sorted $\ell_1$ norm and its dual norm

**Definition 1** The sorted  $\ell_1$  norm associated to  $\lambda \in \mathbb{R}^p$  with  $\lambda_1 \geq \dots \geq \lambda_p \geq 0$  and  $\lambda_1 > 0$  is defined as follows:

$$J_\lambda(b) = \sum_{i=1}^p \lambda_i |b|_{\downarrow i}, \quad b \in \mathbb{R}^p,$$

where  $|b|_{\downarrow 1} \geq \dots \geq |b|_{\downarrow p}$  are the sorted components of  $b$  with respect to the absolute value.

Given a norm  $\|\cdot\|$  on  $\mathbb{R}^p$ , we recall that its dual norm  $\|\cdot\|^*$  is defined by  $\|v\|^* = \max\{b'v : \|b\| \leq 1\}$ , for  $v \in \mathbb{R}^p$ .

**Remark 1** The dual sorted  $\ell_1$  norm has an explicit expression given in [11] and reminded hereafter:

$$J_\lambda^*(v) = \max \left\{ \frac{\|v\|_{(1)}}{\lambda_1}, \frac{\|v\|_{(2)}}{\sum_{i=1}^2 \lambda_i}, \dots, \frac{\|v\|_{(p)}}{\sum_{i=1}^p \lambda_i} \right\}, \quad v \in \mathbb{R}^p,$$

where  $\|\cdot\|_{(k)}$  is the  $k$ -norm (the sum of the  $k$  largest components in absolute value).

### 2.2 SLOPE pattern

The SLOPE pattern introduced in [14], whose definition is reminded below, is a central notion in this article.

**Definition 2** The SLOPE pattern  $\text{patt}(b) \in \mathbb{Z}^p$  of  $b \in \mathbb{R}^p$  is defined by

$$\text{patt}(b)_i = \text{sign}(b_i) \text{rank}(|b|)_i, \quad i \in \{1, \dots, p\},$$

where  $\text{rank}(|b|)_i \in \{0, 1, \dots, k\}$ ,  $k$  is the number of nonzero distinct values in  $\{|b_1|, \dots, |b_p|\}$ ,  $\text{rank}(|b|)_i = 0$  if and only if  $b_i = 0$ , and  $\text{rank}(|b|)_i < \text{rank}(|b|)_j$  if  $|b_i| < |b_j|$ .

We denote by  $\mathcal{P}_p^{\text{slope}} = \text{patt}(\mathbb{R}^p)$  the set of SLOPE patterns. Note in the definition above that  $k = \|\text{patt}(b)\|_\infty$ ; it is the number of nonzero clusters of  $b$ .

**Example 1** Let  $b = (4.2, -1.3, 0, 1.3, 4.2)'$ . Then  $\text{patt}(b) = (2, -1, 0, 1, 2)'$ .

**Definition 3** Let  $m \in \mathbb{Z}^p$  be a SLOPE pattern with  $k := \|m\|_\infty > 0$ . The associated pattern matrix  $U_m \in \mathbb{R}^{p \times k}$  is defined by

$$(U_m)_{ij} = \text{sign}(m_i) \mathbf{1}_{(|m_i|=k+1-j)}, \quad i \in \{1, \dots, p\}, j \in \{1, \dots, k\}.$$

For  $k \geq 1$  we denote  $\mathbb{R}^{k+} = \{s \in \mathbb{R}^k : s_1 > \dots > s_k > 0\}$ . Definition 3 is such that, for  $b \in \mathbb{R}^p$  and  $m \in \mathbb{Z}^p$  a SLOPE pattern with  $k := \|m\|_\infty > 0$ , we have

$$\text{patt}(b) = m \iff \exists s \in \mathbb{R}^{k+} \text{ such that } b = U_m s.$$

Hereafter, the notation  $|m|_{\downarrow} = (|m|_{\downarrow 1}, \dots, |m|_{\downarrow p})'$  represents the components of  $m$  sorted non-increasingly with respect to the absolute value.

<sup>2</sup>This question has been addressed recently in two preprints [12, 18] where, contrary to us, it is required that  $\ker(X) = \{0\}$  and the affine components are not characterized.

**Example 2** Let  $m = (2, -1, 0, 1, 2)'$ . Then

$$U_m = \begin{pmatrix} 1 & 0 & 0 & 0 & 1 \\ 0 & -1 & 0 & 1 & 0 \end{pmatrix}' \text{ and } U_{|m|_{\downarrow}} = \begin{pmatrix} 1 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 1 & 0 \end{pmatrix}'.$$

**Definition 4** Let  $m \in \mathbb{Z}^p$  be a SLOPE pattern with  $k := \|m\|_{\infty} > 0$ . The clustered matrix  $\tilde{X}_m \in \mathbb{R}^{n \times k}$  of  $X \in \mathbb{R}^{n \times p}$  is defined by  $\tilde{X}_m = XU_m$ ; the clustered parameter  $\tilde{\lambda}_m \in \mathbb{R}^k$  of  $\lambda \in \mathbb{R}^p$  is defined by  $\tilde{\lambda}_m = (U_{|m|_{\downarrow}})' \lambda$ .

Note that the dimension of the design matrix  $X$  is reduced when it is clustered as  $\tilde{X}_m$  by a pattern  $m$ : a null component  $m_i = 0$  leads to discard the column  $X_i$  from the design matrix  $X$ , and a cluster  $K \subset \{1, \dots, p\}$  of  $m$  (set of components of  $m$  equal in absolute value) leads to replace the columns  $(X_i)_{i \in K}$  by one column equal to the signed sum:  $\sum_{i \in K} \text{sign}(m_i) X_i$ .

**Example 3** Let  $X = (X_1|X_2|X_3|X_4|X_5)$ ,  $m = (2, -1, 0, 1, 2)'$ ,  $\lambda = (\lambda_1, \lambda_2, \lambda_3, \lambda_4, \lambda_5)' \in \mathbb{R}^5$ . Then the clustered matrix and the clustered parameter are given by:

$$\tilde{X}_m = (X_1 + X_5 | -X_2 + X_4) \text{ and } \tilde{\lambda}_m = \begin{pmatrix} \lambda_1 + \lambda_2 \\ \lambda_3 + \lambda_4 \end{pmatrix}.$$

### 2.3 Subdifferential of the sorted $\ell_1$ norm

The subdifferential of a norm is related to the dual norm  $\|\cdot\|_*$  via the following formula [7, p. 180]:

$$\partial \|\cdot\|(b) = \{v \in \mathbb{R}^p : \|v\|_* \leq 1 \text{ and } b'v = \|b\|\}, \quad b \in \mathbb{R}^p.$$

In particular,  $\partial \|\cdot\|(b)$  is a face of the dual unit ball. For the sorted  $\ell_1$  norm, the above formula can be specified further with the pattern matrix and the clustered parameter associated to  $m = \text{patt}(b)$  for  $b \neq 0$  [2, 14]:

$$\partial J_{\lambda}(b) = \left\{ v \in \mathbb{R}^p : J_{\lambda}^*(v) \leq 1 \text{ and } U_m' v = \tilde{\lambda}_m \right\}. \quad (2)$$

**Remark 2** Let  $\lambda \in \mathbb{R}^{p+}$ . The mapping  $m \mapsto \partial J_{\lambda}(m)$  is a bijection between the set of SLOPE patterns and the set of faces of the unit ball of  $J_{\lambda}^*$  (the signed permutahedron) [14, Theorem 6]. It is no longer true when  $\lambda_1 \geq \dots \geq \lambda_p \geq 0$  is not a decreasing sequence. Therefore we restrict our study to the case where  $\lambda \in \mathbb{R}^{p+}$ , i.e.  $\lambda_1 > \dots > \lambda_p > 0$ . The bijection is illustrated in Figure 1:

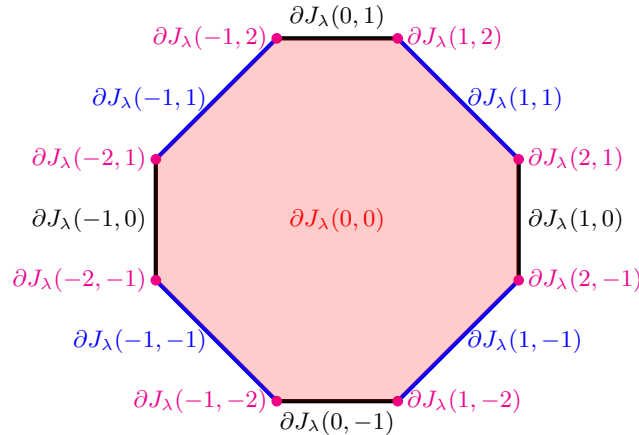


Figure 1: Bijection between

$$\mathcal{P}_2^{\text{slope}} = \{(0, 0), \pm(1, 0), \pm(0, 1), \pm(1, 1), \pm(1, -1), \pm(1, 2), \pm(1, -2), \pm(2, 1), \pm(2, -1)\}$$

and the set of faces of the signed permutahedron for  $\lambda_1 > \lambda_2 > 0$ .

### 3 Solution, fitted value and gradient paths

#### 3.1 Solution set and fitted value

Given  $X \in \mathbb{R}^{n \times p}$ ,  $y \in \mathbb{R}^n$ ,  $\lambda \in \mathbb{R}^{p+}$ , and  $\gamma > 0$ , we denote by  $\mathcal{S}_{X,y,\lambda}(\gamma)$  (or simply  $\mathcal{S}(\gamma)$  when there is no ambiguity) the set of solutions to the SLOPE optimization problem (1), namely:

$$\min_{b \in \mathbb{R}^p} \frac{1}{2} \|y - Xb\|_2^2 + \gamma J_\lambda(b).$$

For any  $\gamma > 0$ , the objective function of the above problem is continuous and coercive thus the solution set  $\mathcal{S}(\gamma)$  is nonempty. Moreover, the fitted value  $\widehat{\text{fit}}(\gamma) = X\widehat{\beta}$  does not depend on  $\widehat{\beta} \in \mathcal{S}(\gamma)$ . When  $\mathcal{S}(\gamma)$  is a singleton, we denote by  $\widehat{\beta}(\gamma)$  its unique element. Note that uniqueness is rather a weak assumption, indeed the set  $\{X \in \mathbb{R}^{n \times p} : \exists y \in \mathbb{R}^n, \exists \gamma > 0 \text{ such that } \mathcal{S}_{X,y,\lambda}(\gamma) \text{ is not a singleton}\}$  has zero Lebesgue measure [14, Proposition 3].

Theorem 1 below shows that  $\widehat{\text{fit}}(\cdot)$  and  $\widehat{\beta}(\cdot)$  are piecewise linear functions. Moreover expressions of  $\widehat{\text{fit}}(\cdot)$  and  $\widehat{\beta}(\cdot)$  restricted to the interval  $I_m$  (depending on a SLOPE pattern  $m$ ) are affine and explicit. We denote hereafter by  $A^+$  the Moore-Penrose pseudo-inverse of a matrix  $A$ .

**Theorem 1** *Let  $X \in \mathbb{R}^{n \times p}$ ,  $y \in \mathbb{R}^n$ ,  $\lambda \in \mathbb{R}^{p+}$ , and  $m \in \mathbb{Z}^p$  be a non-null SLOPE pattern with  $k := \|m\|_\infty > 0$ .*

1. *The set  $I_m := \{\gamma > 0 : \exists \widehat{\beta} \in \mathcal{S}(\gamma) \text{ such that } \text{patt}(\widehat{\beta}) = m\}$  is an interval, with the following characterization:*

$$\begin{aligned} & \gamma \in I_m \\ & \Updownarrow \\ & \begin{cases} \exists s \in \mathbb{R}^{k+} \text{ such that } \tilde{X}'_m y - \gamma \tilde{\lambda}_m = \tilde{X}'_m \tilde{X}_m s & (\text{positivity condition}), \\ X'(\tilde{X}'_m)^+ \tilde{\lambda}_m + \frac{1}{\gamma} X'(I_n - (\tilde{X}'_m)^+ \tilde{X}'_m) y \in \partial J_\lambda(m) & (\text{subdifferential condition}). \end{cases} \end{aligned}$$

*Moreover,  $\widehat{\beta} := U_m s \in \mathcal{S}(\gamma)$  and  $\text{patt}(\widehat{\beta}) = m$  for any  $s \in \mathbb{R}^{k+}$  satisfying the positivity condition at  $\gamma \in I_m$ .*

2. *The fitted value path  $\gamma \mapsto \widehat{\text{fit}}(\gamma)$  is continuous and piecewise linear on  $(0, +\infty)$ , with the following affine expression on  $I_m$ :*

$$\widehat{\text{fit}}(\gamma) = (\tilde{X}'_m)^+ \tilde{X}'_m y - \gamma (\tilde{X}'_m)^+ \tilde{\lambda}_m, \quad \gamma \in I_m.$$

3. *If  $\mathcal{S}(\gamma) = \{\widehat{\beta}(\gamma)\}$  for all  $\gamma > 0$ , then the solution path  $\gamma \mapsto \widehat{\beta}(\gamma)$  is continuous and piecewise linear on  $(0, +\infty)$ , with the following affine expression on  $I_m$ :*

$$\widehat{\beta}(\gamma) = U_m (\tilde{X}'_m \tilde{X}_m)^{-1} (\tilde{X}'_m y - \gamma \tilde{\lambda}_m), \quad \gamma \in I_m.$$

The characterization of the interval  $I_m$  above is closely related to Theorem 3.1 in [2].

#### 3.2 Gradient path and clusters

A solution of the SLOPE optimization problem is characterized by the following two conditions

$$\widehat{\beta} \in \mathcal{S}(\gamma) \Leftrightarrow \begin{cases} J_\lambda^*(X'(y - X\widehat{\beta})) \leq \gamma \\ \widehat{\beta}' X'(y - X\widehat{\beta}) = \gamma J_\lambda(\widehat{\beta}) \end{cases}$$

Note that  $X'(y - X\widehat{\beta}) = X'(y - \widehat{\text{fit}}(\gamma))$  is the gradient at  $\widehat{\beta}$  of the sum of residual squares  $b \mapsto \frac{1}{2} \|y - Xb\|_2^2$ . Subsequently, we call gradient path the expression  $\gamma > 0 \mapsto X'(y - \widehat{\text{fit}}(\gamma))$ . The set of inequalities describing the ball of radius  $\gamma$  for the dual sorted  $\ell_1$  norm which are saturated by the gradient is :

$$\mathcal{A}(\gamma) := \left\{ i \in [p] : \frac{\|X'(y - \widehat{\text{fit}}(\gamma))\|_{(i)}}{\sum_{j=1}^i \lambda_j} = \gamma \right\}.$$

According to Proposition 1 below, the set  $\mathcal{A}(\gamma)$  provides both the number of non-zero clusters, the size of these clusters as well as the number of non-zero components.

**Proposition 1** Let  $\lambda \in \mathbb{R}^{p+}$ ,  $X \in \mathbb{R}^{n \times p}$ ,  $y \in \mathbb{R}^n$ ,  $\gamma > 0$  and  $\widehat{\beta} \in \mathcal{S}(\gamma)$ .

1. Let  $1 \leq k_1 \leq \dots \leq k_l \leq p$  be a subdivision such that:

$$|\text{supp}(\widehat{\beta})| = k_l \text{ and } |\widehat{\beta}|_{\downarrow 1} = \dots = |\widehat{\beta}|_{\downarrow k_1} > \dots > |\widehat{\beta}|_{\downarrow k_{l-1}+1} = \dots = |\widehat{\beta}|_{\downarrow k_l} > 0$$

(i.e.  $\widehat{\beta}$  has  $l$  non-null clusters, the cluster of the largest value has  $k_1$  elements and so on and  $\widehat{\beta}$  has  $k_l$  non-null components). Then,  $\{k_1, \dots, k_l\} \subset \mathcal{A}(\gamma)$ .

2. Conversely, if  $\{k_1, \dots, k_l\} = \mathcal{A}(\gamma)$  then

$$|\widehat{\beta}|_{\downarrow 1} = \dots = |\widehat{\beta}|_{\downarrow k_1} \geq \dots \geq |\widehat{\beta}|_{\downarrow k_{l-1}+1} = \dots = |\widehat{\beta}|_{\downarrow k_l} \geq |\widehat{\beta}|_{\downarrow k_l+1} = \dots = |\widehat{\beta}|_{\downarrow p} = 0$$

(i.e. the number of non-null clusters of  $\widehat{\beta}$  is smaller or equal to  $l$  and the number of non-null components is smaller or equal to  $k_l$ ).

## 4 Algorithms to compute the solution path

To keep this section simple we assume that  $\mathcal{S}(\gamma) = \{\widehat{\beta}(\gamma)\}$  for all  $\gamma > 0$ . Let  $J_\lambda^*(X'y) = \gamma_0 > \gamma_1 > \dots > \gamma_r > \gamma_{r+1} = 0$  be a subdivision such that  $\gamma \mapsto \widehat{\beta}(\gamma)$  is affine with pattern  $m^{(i)}$  on the interval  $(\gamma_{i+1}, \gamma_i)$  for  $i = 0, \dots, r$  (i.e. the interior of  $I_{m^{(i)}}$  is  $(\gamma_{i+1}, \gamma_i)$ ).

First, let us explain how to compute the SLOPE solution path on  $[\gamma_1, \gamma_0]$ . By construction of  $m^{(0)}$  the following implication holds

$$\forall \gamma \in (\gamma_1, \gamma_0), \text{patt}(\widehat{\beta}(\gamma)) = m^{(0)} \Rightarrow \frac{1}{\gamma} X'(y - \widehat{\text{fit}}(\gamma)) \in \partial J_\lambda(m^{(0)}).$$

Moreover, since  $\gamma > 0 \mapsto \widehat{\text{fit}}(\gamma)$  is continuous,  $\widehat{\text{fit}}(\gamma_0) = 0$  and  $\partial J_\lambda(m^{(0)})$  is a closed set, we get

$$\frac{1}{\gamma_0} X'(y - \widehat{\text{fit}}(\gamma_0)) = \frac{1}{\gamma_0} X'y \in \partial J_\lambda(m^{(0)}). \quad (3)$$

Algorithm 1 provides the pattern  $M(\frac{1}{\gamma_0} X'y)$  of the smallest face of the signed permutahedron containing  $\frac{1}{\gamma_0} X'y$ . Therefore, by construction

$$\partial J_\lambda \left( M \left( \frac{1}{\gamma_0} X'y \right) \right) \subset \partial J_\lambda(m^{(0)}) \Rightarrow \left\| M \left( \frac{1}{\gamma_0} X'y \right) \right\|_\infty \leq \|m^{(0)}\|_\infty \quad (4)$$

According to (4), if  $\frac{X'y}{\gamma_0}$  lies onto a facet of the signed permutahedron, we have  $m^{(0)} = M \left( \frac{1}{\gamma_0} X'y \right)$ .

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**Algorithm 1** Pattern of the smallest face containing a vector

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**Require:**  $\lambda \in \mathbb{R}^{p+}$  and  $z \in \mathbb{R}^p$  such that  $J_\lambda^*(z) \leq 1$

Define the set of saturated inequalities as follows

$$\mathcal{A}(z) := \left\{ i \in [p] : \frac{\|z\|^{(i)}}{\sum_{j=1}^i \lambda_j} = 1 \right\}.$$

Define  $M(z) \in \mathcal{P}_p^{\text{slope}}$  as follows

$$M(z) := \begin{cases} 0 & \text{if } \mathcal{A}(z) = \emptyset \\ \forall j \in [p], M_j(z) = \text{sign}(z_j) \sum_{i \in \mathcal{A}(z)} \mathbf{1}(|z_j| \geq \lambda_i) & \text{if } \mathcal{A}(z) \neq \emptyset \end{cases}$$

**return**  $M(z)$

---

**Example 4** We provide illustrations of the solution path of SLOPE when  $y = (6, 2)' \in \mathbb{R}^2$ ,  $\lambda = (4, 2) \in \mathbb{R}^{2+}$  and  $X \in \mathbb{R}^{2 \times 2}$  is the matrix given below

$$X = \begin{pmatrix} 1 & 0.5 \\ 0.5 & 1 \end{pmatrix}.$$

**Largest kink  $\gamma_0$ :** We have  $X'y = (7, 5)'$ , therefore  $\gamma_0 = J_\lambda^*(X'y) = 2$ .

**Pattern  $m^{(0)}$  in the left neighborhood of  $\gamma_0$ :** Since  $\frac{1}{\gamma_0}X'y = (3.5, 2.5)'$  lies in the relative interior of the permutahedron  $\partial J_\lambda(1, 1) = \text{conv}\{(4, 2)', (2, 4)'\}$  then  $m^{(0)} = M(\frac{1}{\gamma_0}X'y) = (1, 1)$ .

**Affine expression of  $\hat{\beta}(\gamma)$  in the left neighborhood of  $\gamma_0$ :** According to statement 3 in Theorem 1 when  $\gamma < \gamma_0 = 2$  is sufficiently close to  $\gamma_0$  we have  $\hat{\beta}(\gamma) = (\frac{8-4\gamma}{3}, \frac{8-4\gamma}{3})$ .

Algorithm 2 uses the characterisation of  $I_{m^{(0)}}$ , based on the positivity and subdifferential conditions, to provide both the kink  $\gamma_1$  as well as the pattern  $m^{(1)}$ .

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**Algorithm 2** Computing the next kink and the next pattern

---

**Require:**  $X \in \mathbb{R}^{n \times p}$ ,  $\lambda \in \mathbb{R}^{p+}$ ,  $\gamma_i, m^{(i)}$  and  $s(\gamma) = (\tilde{X}'_{m^{(i)}} \tilde{X}_{m^{(i)}})^{-1}(\tilde{X}'_{m^{(i)}} y - \gamma \tilde{\lambda}_{m^{(i)}})$

Let  $k = \|m^{(i)}\|_\infty$  and compute

$$\gamma_{\text{fuse}} = \begin{cases} \sup\{\gamma \in [0, \gamma_i) : s(\gamma) \notin \mathbb{R}^{k+}\} & \text{if the set is not empty} \\ 0 & \text{otherwise} \end{cases}$$

**if  $\gamma_{\text{fuse}} = 0$  then**

  Compute  $\gamma_{\text{split}}$  as follows

$$\gamma_{\text{split}} = \begin{cases} \sup\{\gamma \in [0, \gamma_i) : X'(y - \tilde{X}_{m^{(i)}} s(\gamma)) \notin \gamma \partial J_\lambda(m^{(i)})\} & \text{if the set is not empty} \\ 0 & \text{otherwise} \end{cases}$$

**if  $\gamma_{\text{split}} = 0$  then**

**return** The SLOPE solution path is entirely computed

**else**

$$\gamma_{i+1} = \gamma_{\text{split}}$$

  Using Algorithm 1 compute  $m^{(i+1)} = M(\frac{1}{\gamma_{i+1}}X'(y - \tilde{X}_{m^{(i)}} s(\gamma_{i+1})))$

**return**  $\gamma_{i+1}, m^{(i+1)}$

**end if**

**else if  $\gamma_{\text{fuse}} > 0$  and  $X'(y - \tilde{X}_{m^{(i)}} s(\gamma_{\text{fuse}})) \in \gamma_{\text{fuse}} \partial J_\lambda(M)$  then**

$$\gamma_{i+1} = \gamma_{\text{fuse}}$$

$$m^{(i+1)} = \text{patt}(U_{m^{(i)}} s(\gamma_{\text{fuse}}))$$

**return**  $\gamma_{i+1}, m^{(i+1)}$

**else**

  Compute  $\gamma_{i+1}$  as follows

$$\gamma_{i+1} = \sup\{\gamma \in [\gamma_{\text{fuse}}, \gamma_i) : X'(y - \tilde{X}_{m^{(i)}} s(\gamma)) \notin \gamma \partial J_\lambda(m^{(i)})\}$$

  Using Algorithm 1 compute  $m^{(i+1)} = M(\frac{1}{\gamma_{i+1}}X'(y - \tilde{X}_{m^{(i)}} s(\gamma_{i+1})))$

**return**  $\gamma_{i+1}, m^{(i+1)}$

**end if**

---

Using iteratively Algorithm 2 allows to compute entirely the SLOPE solution path<sup>3</sup>.

## 5 SLOPE solution path applied on data

Below we will use a data set describing the quality of red ‘‘Vinho Verde’’ wines [4]<sup>4</sup>. In this data set explanatory variables  $X \in \mathbb{R}^{1599 \times 11}$  are physicochemical measurements such as density, acidity, the amount of sugar and alcohol, etc and the response  $y \in \mathbb{R}^{1599}$  is the wine quality score between 0

<sup>3</sup>An implementation in Python of these algorithms for computing the SLOPE solution path is available online: <https://anonymous.4open.science/r/slope-path-744C>

<sup>4</sup>This data set is available online: <http://archive.ics.uci.edu/ml/datasets/Wine+Quality>

and 10. For this numerical experiment we take  $\lambda = (11, 10, \dots, 1)$ . The matrix  $X$  is mean-centered ( $\forall j \in \{1, \dots, 11\} \sum_{i=1}^{1599} X_{ij} = 0$ ) and standardized ( $\forall j \in \{1, \dots, 11\} \sum_{i=1}^{1599} X_{ij}^2 = 1598$ ).

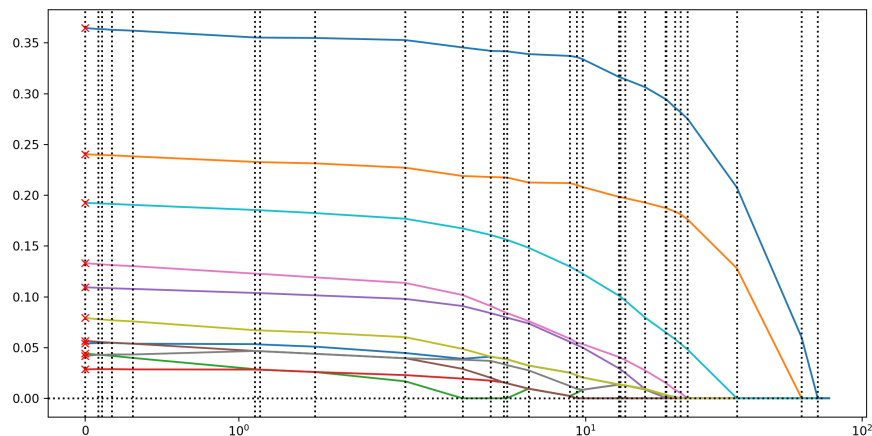


Figure 2: This figure provides the solution of SLOPE, in absolute value, as a function of  $\gamma > 0$  (the x-axis is reported on the logarithm scale). One may observe that when  $0.07 = \gamma^{(28)} < \gamma < \gamma^{(4)} = 17.79$ , at least two components of SLOPE are equal in absolute value. Moreover, in absolute value, the SLOPE solution converges to the ordinary least squares estimator when  $\gamma$  tends to 0 (components of the ordinary least squares estimator in absolute value are crosses on the y-axis).

## 6 Conclusion and future works

One of the main result in this article is Theorem 1 proving that the SLOPE solution path is piecewise linear and providing the characterization of the intervals where the path is affine. Moreover algorithms 1 and 2 allow to compute numerically this path. The computational time of our numerical scheme depends mainly on the number of kinks. In the illustration on real data set, the number of intervals is rather small (29) and kinks are concentrated around zero. However the number of intervals where the path is affine is bounded by the number of SLOPE patterns in  $\mathbb{R}^p$  and potentially, similarly as for LASSO [9], this huge upper bound might be reached for some pathological examples. In the future we are going to test our method on various data sets and if kinks are too concentrated around zero we would switch for another numerical scheme (like, for instance, the one developed in [8]) to compute the SLOPE solution path in the neighborhood of zero.

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## 7 Appendix

### Proof of Theorem 1

**1:  $I_m$  is an interval**) Hereafter we suppose that  $I_m \neq \emptyset$ . Let  $\gamma_0, \gamma_1 \in I_m$  and pick  $\hat{\beta}(\gamma_0) \in \mathcal{S}(\gamma_0)$ ,  $\hat{\beta}(\gamma_1) \in \mathcal{S}(\gamma_1)$  such that  $\text{patt}(\hat{\beta}(\gamma_0)) = \text{patt}(\hat{\beta}(\gamma_1)) = m$ . Let  $\alpha \in [0, 1]$ ,  $\bar{\gamma} = \alpha\gamma_0 + (1 - \alpha)\gamma_1$  and  $\bar{\beta} = \alpha\hat{\beta}(\gamma_0) + (1 - \alpha)\hat{\beta}(\gamma_1)$  then  $\text{patt}(\bar{\beta}) = m$ . Indeed, if  $m = 0$  then clearly  $\text{patt}(\bar{\beta}) = 0$ . Otherwise, let  $k = \|m\|_\infty \geq 1$  then  $\hat{\beta}(\gamma_0) = U_m s_0$  for some  $s_0 \in \mathbb{R}^{k+}$ ,  $\hat{\beta}(\gamma_1) = U_m s_1$  for some



$s_1 \in \mathbb{R}^{k+}$  therefore  $\bar{\beta} = U_m \bar{s}$  where  $\bar{s} = \alpha s_0 + (1 - \alpha) s_1 \in \mathbb{R}^{k+}$ . To prove that  $I_m$  is an interval it remains to show that  $\bar{\beta} \in \mathcal{S}(\bar{\gamma})$ . Because both  $\hat{\beta}(\gamma_0)$  and  $\hat{\beta}(\gamma_1)$  are SLOPE minimizers, we have

$$X'(y - X\hat{\beta}(\gamma_0)) \in \gamma_0 \partial J_\lambda(m) \text{ and } X'(y - X\hat{\beta}(\gamma_1)) \in \gamma_1 \partial J_\lambda(m).$$

By construction of  $\bar{\beta}$  the following equality occurs:

$$\alpha X'(y - X\hat{\beta}(\gamma_0)) + (1 - \alpha) X'(y - X\hat{\beta}(\gamma_1)) = X'(y - X\bar{\beta}).$$

Moreover, since  $\partial J_\lambda(m)$  is a convex set, we have  $\alpha \gamma_0 \partial J_\lambda(m) + (1 - \alpha) \gamma_1 \partial J_\lambda(m) \subset \bar{\gamma} \partial J_\lambda(m)$ . Consequently,  $X'(y - X\bar{\beta}) \in \bar{\gamma} \partial J_\lambda(m) = \bar{\gamma} \partial J_\lambda(\bar{\beta})$  thus  $\bar{\beta} \in \mathcal{S}(\bar{\gamma})$ .

**1: characterization of  $I_m$**  The proof of this characterization is closely related to the proof of Theorem 3.1 in [2].

*Necessity.* If  $\gamma \in I_m$ , then there exists  $\hat{\beta} \in \mathcal{S}(\gamma)$  such that  $\text{patt}(\hat{\beta}) = m$ . Consequently,  $\hat{\beta} = U_m s$  for some  $s \in \mathbb{R}^{k+}$ . Because  $\hat{\beta}$  is a element of  $\mathcal{S}(\gamma)$  whose pattern is  $m$  then  $X'(y - \hat{\text{fit}}(\gamma)) \in \gamma \partial J_\lambda(\hat{\beta}) = \gamma \partial J_\lambda(m)$ . Multiplying this inclusion by  $U'_m$ , due to (2), we get  $\tilde{X}'_m(y - \hat{\text{fit}}(\gamma)) = \gamma \tilde{\lambda}_m$  and so

$$\tilde{X}'_m y - \gamma \tilde{\lambda}_m = \tilde{X}'_m \hat{\text{fit}}(\gamma) = \tilde{X}'_m \tilde{X}_m s. \quad (5)$$

The positivity condition is proven.

We apply  $(\tilde{X}'_m)^+$  from the left to (5) and use the fact that  $(\tilde{X}'_m)^+ \tilde{X}'_m$  is the projection onto  $\text{col}(\tilde{X}_m)$ . Since  $\hat{\text{fit}}(\gamma) \in \text{col}(\tilde{X}_m)$ , we have  $(\tilde{X}'_m)^+ \tilde{X}'_m \hat{\text{fit}}(\gamma) = \hat{\text{fit}}(\gamma)$ . Thus,

$$(\tilde{X}'_m)^+ \tilde{X}'_m y - \gamma (\tilde{X}'_m)^+ \tilde{\lambda}_m = \hat{\text{fit}}(\gamma).$$

The above equality gives the subdifferential condition:

$$\begin{aligned} \partial J_\lambda(m) \ni \frac{1}{\gamma} X'(y - \hat{\text{fit}}(\gamma)) &= \frac{1}{\gamma} X'(y - ((\tilde{X}'_m)^+ \tilde{X}'_m y - \gamma (\tilde{X}'_m)^+ \tilde{\lambda}_m)) \\ &= X'(\tilde{X}'_m)^+ \tilde{\lambda}_m + \frac{1}{\gamma} X'(I_n - (\tilde{X}'_m)^+ \tilde{X}'_m) y. \end{aligned} \quad (6)$$

*Sufficiency.* Assume that the positivity condition and the subdifferential conditions hold true. Then, by the positivity condition, one may pick  $s \in \mathbb{R}^{k+}$  for which

$$\gamma \tilde{\lambda}_m = \tilde{X}'_m y - \tilde{X}'_m \tilde{X}_m s. \quad (7)$$

Let us show that  $U_m s \in \mathcal{S}(\gamma)$ . By definition of  $U_m$ , we have  $\text{patt}(U_m s) = m$  thus  $\partial J_\lambda(U_m s(\gamma)) = \partial J_\lambda(m)$ . Moreover, using (6) and (7) one may deduce

$$\begin{aligned} \partial J_\lambda(U_m s) &\ni \frac{1}{\gamma} X'(y - (\tilde{X}'_m)^+ \tilde{X}'_m y + \gamma (\tilde{X}'_m)^+ \tilde{\lambda}_m) \\ &= \frac{1}{\gamma} X'(y - (\tilde{X}'_m)^+ \tilde{X}'_m y + (\tilde{X}'_m)^+ (\tilde{X}_m y - \tilde{X}'_m \tilde{X}_m s)) \\ &= \frac{1}{\gamma} X'(y - XU_m s). \end{aligned}$$

Consequently  $U_m s \in \mathcal{S}(\gamma)$ .

**2 and 3: continuity** Let  $\gamma \in (0, +\infty)$ ,  $(\gamma_n)_{n \in \mathbb{N}}$  be a sequence converging to  $\gamma$  and  $\hat{\beta}(\gamma_n) \in S_{X, \gamma_n J_\lambda}(y)$ . Both sequences  $(\hat{\beta}(\gamma_n))_{n \in \mathbb{N}}$  and  $(\hat{\text{fit}}(\gamma_n))_{n \in \mathbb{N}}$  are bounded therefore, up to extract a subsequence, one may assume that both  $(\hat{\beta}(\gamma_n))_{n \in \mathbb{N}}$  and  $(\hat{\text{fit}}(\gamma_n))_{n \in \mathbb{N}}$  converge respectively to a limit point  $l \in \mathbb{R}^p$  and  $Xl \in \mathbb{R}^n$ . Let  $\hat{\beta}(\gamma) \in \mathcal{S}(\gamma)$ . Because  $\hat{\beta}(\gamma_n)$  is a minimizer, the following inequality occurs.

$$\frac{1}{2} \|y - \hat{\text{fit}}(\gamma_n)\|_2^2 + \gamma_n J_\lambda(\hat{\beta}(\gamma_n)) \leq \frac{1}{2} \|y - \hat{\text{fit}}(\gamma)\|_2^2 + \gamma_n J_\lambda(\hat{\beta}(\gamma)).$$

Taking the limit in the above expression gives

$$\frac{1}{2} \|y - Xl\|_2^2 + \gamma J_\lambda(l) \leq \frac{1}{2} \|y - \hat{\text{fit}}(\gamma)\|_2^2 + \gamma J_\lambda(\hat{\beta}(\gamma)).$$

Because  $\widehat{\beta}(\gamma) \in \mathcal{S}(\gamma)$ , one may deduce that  $l \in \mathcal{S}(\gamma)$  and thus  $Xl = \widehat{\text{fit}}(\gamma)$ . Therefore, the unique limit point of the bounded sequence  $(\widehat{\text{fit}}(\gamma_n))_{n \in \mathbb{N}}$  is  $\widehat{\text{fit}}(\gamma)$ . Consequently,  $\lim_{n \rightarrow +\infty} \widehat{\text{fit}}(\gamma_n) = \widehat{\text{fit}}(\gamma)$  and thus the function  $\gamma \in (0, +\infty) \mapsto \widehat{\text{fit}}(\gamma)$  is continuous. Similarly, if  $\mathcal{S}(\gamma)$  is a singleton then  $l = \widehat{\beta}(\gamma)$ , the unique limit point of the bounded sequence  $(\widehat{\beta}(\gamma_n))_{n \in \mathbb{N}}$  is  $\widehat{\beta}(\gamma)$  and thus  $\lim_{n \rightarrow +\infty} \widehat{\beta}(\gamma_n) = \widehat{\beta}(\gamma)$ . Therefore the function  $\gamma \in (0, +\infty) \mapsto \widehat{\beta}(\gamma)$  is continuous.

2) When  $\gamma \in I_m$  then multiplying both side of the positivity condition by  $(\tilde{X}'_m)^+$  and using the fact that  $(\tilde{X}'_m)^+ \tilde{X}'_m$  is the projection onto  $\text{col}(\tilde{X}_m)$  gives

$$(\tilde{X}'_m)^+ \tilde{X}'_m y - \gamma (\tilde{X}'_m)^+ \tilde{\lambda}_m = (\tilde{X}'_m)^+ \tilde{X}'_m \tilde{X}_m s = \tilde{X}_m s = \widehat{\text{fit}}(\gamma).$$

3) The proof of statement 3) relies on Lemma 1 proved in supplementary material.

**Lemma 1** *Let  $X \in \mathbb{R}^{n \times p}$ ,  $y \in \mathbb{R}^n$  and  $\lambda \in \mathbb{R}^{p+}$ . There exists  $\widehat{\beta} \in \mathcal{S}(\gamma)$  for which the pattern  $m = \text{patt}(\widehat{\beta})$  satisfies  $\ker(\tilde{X}_m) = \{0\}$ .*

Consequently, when  $\gamma \in I_m$  and  $\mathcal{S}(\gamma)$  is a singleton then  $\ker(\tilde{X}_m) = \{0\}$ , where  $m = \text{patt}(\widehat{\beta}(\gamma))$ . Since  $\tilde{X}'_m \tilde{X}_m$  is invertible, the positivity condition gives

$$\widehat{\beta}(\gamma) = U_m s = U_m (\tilde{X}'_m \tilde{X}_m)^{-1} (\tilde{X}'_m y - \gamma \tilde{\lambda}_m).$$

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