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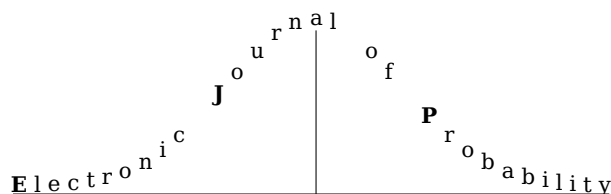
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Non-directed polymers in heavy-tail random environment in dimension $d \geq 2$

Quentin Berger* Niccolò Torri† Ran Wei‡

Abstract

In this article we study a *non-directed* polymer model in dimension $d \geq 2$: we consider a simple symmetric random walk on \mathbb{Z}^d which interacts with a random environment, represented by i.i.d. random variables $(\omega_x)_{x \in \mathbb{Z}^d}$. The model consists in modifying the law of the random walk up to time (or length) N by the exponential of $\sum_{x \in \mathcal{R}_N} \beta(\omega_x - h)$ where \mathcal{R}_N is the range of the walk, i.e. the set of visited sites up to time N , and $\beta \geq 0$, $h \in \mathbb{R}$ are two parameters. We study the behavior of the model in a weak-coupling regime, that is taking $\beta := \beta_N$ vanishing as the length N goes to infinity, and in the case where the random variables ω have a heavy tail with exponent $\alpha \in (0, d)$. We are able to obtain precisely the behavior of polymer trajectories under all possible weak-coupling regimes $\beta_N = \hat{\beta}N^{-\gamma}$ with $\gamma \geq 0$: we find the correct transversal fluctuation exponent ξ for the polymer (it depends on α and γ) and we give the limiting distribution of the rescaled log-partition function. This extends existing works to the non-directed case and to higher dimensions.

Keywords: Random polymer, Random walk, Range, Heavy-tail distributions, Weak-coupling limit, Super-diffusivity, Sub-diffusivity.

MSC2020 subject classifications: 82D60, 60K37, 60G70.

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1 Introduction

We consider in this paper a *non-directed polymer model* in which a simple symmetric random walk on \mathbb{Z}^d (with $d \geq 2$) interacts with a random environment. This model is closely related to the celebrated *directed polymer model*, for which we refer to [17] for a complete overview. The main difference here is that the random walk is *not* directed:

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it can come back to a site that has already been visited, but the interaction is however counted only once (in the spirit of the excited random walk, see [3]). The model can also be viewed as a randomly perturbed version of random walks penalized or rewarded by their ranges, depending on the positivity or negativity of the external field. The non-directed polymer model was first introduced by [25] (with exponential moment conditions on the environment) and its study was followed in [5] with a detailed analysis in dimension $d = 1$. Here, we investigate the case of a heavy-tail random environment, in analogy with [7, 19, 24] for the directed polymer model — our results can be viewed as an extension of [7, 19, 24] to a non-directed setting (and to higher dimensions).

1.1 The non-directed polymer model

Let $S = (S_n)_{n \geq 0}$ be a simple symmetric random walk on \mathbb{Z}^d with $d \geq 1$. We denote the probability and expectation with respect to S by \mathbf{P} and \mathbf{E} respectively. Let $\omega = (\omega_x)_{x \in \mathbb{Z}^d}$ be a field of i.i.d. random variables (the environment), independent of S . The probability and expectation with respect to ω are denoted by \mathbb{P} and \mathbb{E} respectively.

The *non-directed polymer model* up to time N is defined via the following Gibbs' transform: for a given realization ω of the random environment and for $\beta \geq 0$ (the inverse temperature) and $h \in \mathbb{R}$ (an external field), let

$$\frac{d\mathbf{P}_{N,\beta}^{\omega,h}}{d\mathbf{P}}(S) := \frac{1}{Z_{N,\beta}^{\omega,h}} \exp \left(\sum_{x \in \mathcal{R}_N} \beta(\omega_x - h) \right), \quad (1.1)$$

where $\mathcal{R}_N := \{S_0, \dots, S_N\}$ is the range of the random walk up to time N . The partition function $Z_{N,\beta}^{\omega,h}$ is the constant which makes $\mathbf{P}_{N,\beta}^{\omega,h}$ a probability measure and is equal to

$$Z_{N,\beta}^{\omega,h} := \mathbf{E} \left[\exp \left(\sum_{x \in \mathcal{R}_N} \beta(\omega_x - h) \right) \right]. \quad (1.2)$$

Note that if $\omega \equiv 0$ (i.e. when there is no disorder) and $\beta h \equiv c \in (0, +\infty)$, then (1.1) becomes the model of a random walk penalized by its range, defined by $\frac{d\mathbf{P}_{N,c}}{d\mathbf{P}}(S) := \frac{1}{Z_{N,c}} e^{-c|\mathcal{R}_N|}$. That model has been well-studied (starting with the seminal work [21]) and is now well-understood: with high $\mathbf{P}_{N,c}$ -probability, \mathcal{R}_N is close to a d -dimensional ball with an explicit radius $\rho_{d,c} N^{1/(d+2)}$ without holes (see [11, 4, 20]).

When ω is non-trivial, the model (1.1) describes a self-attracting (if $h > 0$) or self-repulsing (if $h < 0$) polymer interacting with a random environment. At each site, the polymer chain interacts with the disorder exactly once, which may model a *screened* interaction, one monomer “absorbing” all the interaction at a specific site. We have chosen to stick to this setting in order to pursue the study initiated in [5, 25], which considers this model as a disordered version of a random walk penalized by its range.

Remark 1.1. One could study a polymer interacting repeatedly with the disorder, that is

$$\frac{d\tilde{\mathbf{P}}_{N,\beta}^{\omega,h}}{d\mathbf{P}}(S) := \frac{1}{\tilde{Z}_{N,\beta}^{\omega,h}} \exp \left(\sum_{i=1}^N \beta(\omega_{S_i} - h) \right)$$

(note that the term $-\beta h N$ in the Hamiltonian is a constant and thus does not change the measure). In that case, large values of ω_x will play an overwhelming role in the measure, since their effect can be accumulated by returning repeatedly to a given site. One can therefore expect a very strong localization phenomenon, the polymer staying on one site with the largest possible ω_x , with high probability. In the literature, this model is referred to as the parabolic Anderson model and the one-site localization for heavy-tailed environment has been proven in [13], see [28] for a continuous-time version of the result.

The main goal is then to understand the shape of typical polymer trajectories $(S_i)_{0 \leq i \leq N}$ under the measure $\mathbf{P}_{N,\beta}^{\omega,h}$, as $N \rightarrow \infty$. One of the main question is to know if and how the presence of the environment perturbs the structure of the typical trajectories of the walk. In particular, one is interested in describing the end-to-end (or wandering) exponent ξ and the fluctuation (or volume) exponent χ , that are loosely speaking defined as:

$$\mathbf{E}_{N,\beta}^{\omega,h}(\|S_N\|^2) \approx N^{2\xi}, \quad \text{Var}(\log Z_{N,\beta}^{\omega,h}) \approx N^{2\chi}.$$

For comparison, let us stress that the random walk $(S_n)_{n \geq 0}$ in dimension d , is replaced in the directed polymer model by a directed random walk $(n, S_n)_{n \geq 0}$ in dimension $1 + d$. In particular, the parameter h does not have any influence on the polymer measure, since the number of visited sites is deterministic (and no site is visited twice). An important feature of the directed polymer model is that a phase transition is known to occur: when ω_x have an exponential moment, if β is (strictly) smaller than some critical value β_c , then trajectories are diffusive ($\xi = 1/2$, $\chi = 0$), whereas if β is (strictly) larger than β_c trajectories exhibit some localization properties and are conjectured to have a super-diffusive behavior (at least in low dimensions) — for instance, in dimension $1 + 1$ it is conjectured that $\xi = 2/3$ and $\chi = 1/3$. The critical value β_c is known to be $\beta_c = 0$ in dimension $1 + 1$ and $1 + 2$ (hence there is no phase transition) and $\beta_c > 0$ in dimension $1 + d$ with $d \geq 3$. We refer to [17] and references therein for more details.

In the non-directed case, however, the parameter h plays an important role, and random walk trajectories with a large range are rewarded or penalized (depending on whether h is positive or negative). In analogy with the directed polymer model, the presence of a random environment should still have a stretching effect. The competition between the folding effect of range penalties and the stretching effect of a random environment has been recently investigated in detail in the case of dimension $d = 1$, [5] (in particular, it has been found that $\xi = 2/3$ when $h = 0$). But such a study appears difficult in higher dimension, because the range of the simple random walk then has a complex geometry. However, in the case of a heavy-tail random environment, the localization features of the model become more salient, since a few sites in the environment will have a much higher value than the others: we are indeed able to describe quite precisely the behavior of the non-directed polymers in that case. This generalizes the study in [7, 19] to the case of non-directed polymers and to higher dimensions. Let us stress that as a by-product of our results we have that in the heavy-tail setting there is no phase transition: $\beta_c = 0$ in any dimension.

Weak-coupling regime.

A recent approach for studying disordered system, initiated in [1], and which has been developed extensively over the past few years, is to consider weak-coupling regimes, see e.g. [6, 14, 15, 19, 25]. The idea is to take the inverse temperature β_N vanishing as N goes to infinity. In the papers cited above, the goal is to find the appropriate scaling for β_N in order for the partition function to converge to a non-trivial random variable. This regime is called of *intermediate disorder*: loosely speaking, it corresponds to a regime in which disorder just *kicks in*, in the sense that it is still felt in the limit, but is not strong enough to make the (averaged) disordered measure singular with respect to the reference measure.

We will assume that there is some $\gamma \geq 0$ and some $\beta \in (0, \infty)$ such that

$$\lim_{N \rightarrow \infty} N^\gamma \beta_N = \beta, \quad \text{as } N \rightarrow +\infty, \quad (1.3)$$

and we will write $\beta_N \sim \beta N^{-\gamma}$; we may also consider the case $\beta = 0$ or $\beta = +\infty$. We think of γ as a parameter one can play with, which tunes the speed at which β_N goes to zero.

We could actually work with general sequences $(\beta_N)_{N \geq 1}$ (in particular, the important relation is (1.5) below), but we stick to the pure power case (1.3) in order to clarify the exposition — it will capture all the essential features of the model, see Remark 2.9.

Let us also stress that we will consider h fixed, seen as a centering term for the disorder ω_x .

Heavy-tail environment.

Our main assumption in this paper is that $\omega = (\omega_x)_{x \in \mathbb{Z}^d}$ are i.i.d. random variables, with a pure power tail behavior. We assume that there exists some $\alpha > 0$ such that

$$\mathbb{P}(\omega_0 > t) \sim t^{-\alpha}, \quad \text{as } t \rightarrow +\infty. \quad (1.4)$$

Also here, we could consider a more general asymptotic behavior in (1.4), for instance replacing the pure power $t^{-\alpha}$ by $L(t)t^{-\alpha}$ with $L(\cdot)$ a slowly varying function. However, we stick to the pure power case as in (1.4) for the sake of clarity — again, it will capture all the essential features of the model.

For simplicity, we also assume that $\omega \geq 0$, but this does not hide anything deep.

1.2 Heuristics for the phase diagram

Let us present a Flory argument to guess the wandering and fluctuation exponents. The idea is to find the correct transversal fluctuations r_N (with $N^{1/2} \leq r_N \leq N$) such that the entropic cost for the random walk to stretch to a distance r_N is balanced by the possible energetic gain from high weights ω_x contained in a ball of radius r_N . At the exponential level:

- the entropic cost is of order r_N^2/N ;
- the energetic gain is of order $\beta_N (r_N^d)^{1/\alpha}$, where $(r_N^d)^{1/\alpha}$ is the order of the maximal weight ω_x in a ball of radius r_N .

Hence, the energy-entropy balance leads to the relation

$$\beta_N (r_N)^{d/\alpha} \sim \frac{r_N^2}{N} \quad (N^{1/2} \leq r_N \leq N). \quad (1.5)$$

In the case we are interested in, that is if $\beta_N \sim N^{-\gamma}$ with $\gamma \geq 0$, this gives $r_N \sim N^\xi$ with ξ verifying

$$\frac{\xi d}{\alpha} - \gamma = 2\xi - 1 \quad \Longleftrightarrow \quad \xi = \frac{\alpha(1-\gamma)}{2\alpha-d}, \quad (1.6)$$

provided that $\xi \in [\frac{1}{2}, 1]$, that is $\frac{d-\alpha}{\alpha} \leq \gamma \leq \frac{d}{2\alpha}$ (note that this range of parameter γ is non-empty only if $\alpha > d/2$). However, this picture should fail when α is large (in particular when $\alpha > d$), because then the strategy of visiting mostly high-energy sites should be outperformed by a *collective* optimization (that is still poorly understood, especially in dimension $d \geq 3$). We refer to [10] for a discussion on that matter. We therefore focus our attention on the case $\alpha < d$.

When $0 \leq \gamma < \frac{d-\alpha}{\alpha}$, then since the transversal fluctuations cannot exceed N , the energetic gain should always overcome the entropic cost: we should have $\xi = 1$. On the other hand, when $\gamma > \frac{d}{2\alpha}$ for $\alpha > d/2$ or $\gamma > \frac{d-\alpha}{\alpha}$ for $\alpha < d/2$, then the energetic gain can never overcome the entropic cost of reaching distance much larger than $N^{1/2}$: we should have $\xi = 1/2$.

To summarize, one should have the following three regions when $\alpha < d$.

- If $\alpha \in (0, d)$ and $\gamma < \frac{d-\alpha}{\alpha}$, then $\xi = 1$.
- If $\alpha \in (\frac{d}{2}, d)$ and $\gamma \in (\frac{d-\alpha}{\alpha}, \frac{d}{2\alpha})$, then $\xi = \frac{\alpha(1-\gamma)}{2\alpha-d} \in (\frac{1}{2}, 1)$.

C. If $\alpha \in (0, d)$ and $\gamma > \max\{\frac{d}{2\alpha}, \frac{d-\alpha}{\alpha}\}$, then $\xi = \frac{1}{2}$.

We refer to Figure 1 below for a graphical representation of these regions, and to [7, Fig. 1] for a comparison with the corresponding picture in the directed polymer in dimension $1 + 1$ (where conjectures for larger values of α have the merit to exist). Our main results consist in establishing this phase diagram.

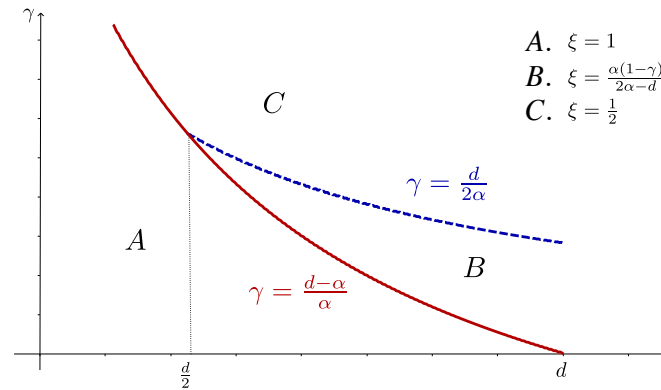


Figure 1: Phase diagram for the wandering exponent ξ depending on the parameters α, γ , with $\alpha \in (0, d)$. We have identified three regions. In region **A**, we have $\xi = 1$. In region **C**, we have $\xi = \frac{1}{2}$. In region **B**, we have $\xi = \frac{\alpha(1-\gamma)}{2\alpha-d} \in (\frac{1}{2}, 1)$. Note that when $\alpha \in (\frac{d}{2}, d)$, the wandering exponent ξ in region **B** interpolates between regions **A** and **C**: all values $\xi \in (\frac{1}{2}, 1)$ can be attained. On the other hand, region **B** does not exist when $\alpha \in (0, \frac{d}{2})$: the wandering exponent ξ drops abruptly from 1 to $\frac{1}{2}$ and no intermediate transversal fluctuations are possible.

1.3 Definition of quantities arising in the scaling limit

Our results additionally provide the scaling limit of the log-partition function in all weak-coupling regimes when $\alpha \in (0, d)$. In order to describe the limit, let us introduce the continuum counterparts of the random environment and of random walk trajectories, and define the corresponding notion of energy and entropy.

The continuum environment arises as the extremal field of $(\omega_x)_{x \in \mathbb{Z}^d}$: we let $\mathcal{P} := \{(x_i, w_i)\}$ be a Poisson point process on $\mathbb{R}^d \times (0, +\infty)$ of intensity $\eta(dx, dw) = \alpha w^{-(1+\alpha)} \mathbb{1}_{\{w>0\}} dx dw$. With a slight abuse of notation (it will not draw any confusion), we denote its law by \mathbb{P} . We also let

$$\mathcal{D} := \{s : [0, 1] \rightarrow \mathbb{R}^d : s(0) = 0, s \text{ continuous and a.e. differentiable}\}, \quad (1.7)$$

which represents the set of allowed paths (corresponding to scaling limits of random walk trajectories). Then, for a path $s \in \mathcal{D}$, we define its energy

$$\pi(s) = \pi^{\mathcal{P}}(s) := \sum_{(x,w) \in \mathcal{P}} w \mathbb{1}_{\{x \in s([0,1])\}}. \quad (1.8)$$

To define the entropic cost, let us observe that we have the following large deviation principles for the simple random walk: for $u \in (0, 1]$ and $x \in \mathbb{R}^d$ (we omit integer parts to lighten notation), cf. Lemma B.1 in Appendix

$$\lim_{N \rightarrow \infty} -\frac{1}{N^{2\xi-1}} \log \mathbf{P}(S_{uN} = xN^\xi) = \begin{cases} u J_d(\frac{x}{u}) & \text{if } \xi = 1, \\ \frac{d}{2} \frac{\|x\|^2}{u} & \text{if } \xi \in (\frac{1}{2}, 1), \end{cases}$$

where $J_d(\cdot)$ is a given rate function; we stress that $J(x) < +\infty$ if $\|x\|_1 \leq 1$ and $J(x) = +\infty$ if $\|x\|_1 > 1$, with $\|x\|_1$ the ℓ^1 norm of x .

We then define two different entropic cost functions for a path $s \in \mathcal{D}$ — or rather for the image $s([0, 1])$ — depending on whether the path comes from the scaling of a random walk trajectory with wandering exponent $\xi = 1$ or $\xi \in (\frac{1}{2}, 1)$. In the case $\xi = 1$, we define

$$\widehat{\text{Ent}}(s) := \inf_{\varphi \in \Phi} \int_0^1 J_d((s \circ \varphi)'(t)) dt, \quad (1.9)$$

where the infimum is over

$$\Phi := \{\varphi : [0, 1] \xrightarrow{\text{onto}} [0, 1] : \varphi \text{ is non-decreasing and a.e. differentiable}\}. \quad (1.10)$$

In the case $\xi \in (\frac{1}{2}, 1)$, we define analogously

$$\text{Ent}(s) = \inf_{\varphi \in \Phi} \int_0^1 \frac{d}{2} \|(s \circ \varphi)'(t)\|^2 dt = \frac{d}{2} \left(\int_0^1 \|s'(t)\| dt \right)^2. \quad (1.11)$$

The second identity comes from the fact that: (i) the right-hand side is a lower bound for the left-hand side (by Cauchy–Schwarz inequality); (ii) the lower bound is attained for the parametrization of s by its length, that is choosing φ such that $\int_0^{\varphi(u)} \|s'(t)\| dt = u \int_0^1 \|s'(t)\| dt$ for $u \in [0, 1]$.

Let us also stress that $J_d(x) \geq \frac{1}{2} \|x\|^2$ for all $x \in \mathbb{R}^d$, so that $\widehat{\text{Ent}}(s) \geq \frac{1}{d} \text{Ent}(s)$ for all $s \in \mathcal{D}$.

The continuous energy-entropy variational problem that we expect to arise as the scaling limit of the log-partition function are $\sup_{s \in \mathcal{D}} \{\beta \pi(s) - \text{Ent}(s)\}$, the entropy term $\text{Ent}(s)$ being one of (1.9) or (1.11) depending on the scaling considered. As an important part of the proofs, we will show that these continuous variational problems are well-defined.

2 Main Results

Before we state our results, let us define more precisely what we intend when we say that the polymer has transversal fluctuations N^ξ .

Definition 2.1. We say that $(S_n)_{0 \leq n \leq N}$ has transversal fluctuations of order r_N under $\mathbf{P}_{N, \beta_N}^{\omega, h}$ if for any $\varepsilon > 0$ there exists some $\eta \in (0, 1)$ such that for all N large enough

$$\mathbf{P}_{N, \beta_N}^{\omega, h} \left(\max_{1 \leq n \leq N} \|S_n\| \in [\eta, \frac{1}{\eta}] r_N \right) \geq 1 - \varepsilon \quad \text{with } \mathbb{P}\text{-probability larger than } 1 - \varepsilon.$$

We separate our results according to the different regions we consider. In all the rest of the paper, we work in dimension $d \geq 2$.

Notational disclaimer. For two sequences $(a_n)_{n \geq 1}, (b_n)_{n \geq 1}$, we write $a_n \sim b_n$ if $\lim_{n \rightarrow \infty} a_n/b_n = 1$, $a_n \ll b_n$ if $\lim_{n \rightarrow \infty} a_n/b_n = 0$, and $a_n \asymp b_n$ if $0 < \liminf a_n/b_n \leq \limsup a_n/b_n < \infty$.

2.1 Region A: $\alpha \in (0, d)$, $\gamma \leq \frac{d-\alpha}{\alpha}$

Our first result gives the scaling limit of the model in region **A** of Figure 1. This is the analogue of what Auffinger and Louidor [2] proved in the context of the directed polymer model in dimension $1 + 1$: the following result therefore generalizes [2] to the non-directed framework and to higher dimensions.

Theorem 2.2. Let $\alpha \in (0, d)$ and let $(\beta_N)_{N \geq 1}$ be such that $\lim_{N \rightarrow \infty} N^{\frac{d-\alpha}{\alpha}} \beta_N = \beta \in (0, +\infty]$. Then for any fixed $h \in \mathbb{R}$, we have the following convergence in distribution, as $N \rightarrow \infty$,

$$\frac{1}{\beta_N N^{d/\alpha}} \log Z_{N, \beta_N}^{\omega, h} \xrightarrow{(d)} \widehat{\mathcal{T}}_\beta := \sup_{s \in \mathcal{D}, \widehat{\text{Ent}}(s) < +\infty} \left\{ \pi(s) - \frac{1}{\beta} \widehat{\text{Ent}}(s) \right\}, \quad (2.1)$$

where $\pi(s)$ and $\widehat{\text{Ent}}(s)$ are defined in (1.8) and (1.9) respectively.

Remark 2.3. Let us stress that by an extended version of Skorokhod representation theorem, see [27, Cor. 5.12], one can upgrade the convergence (2.1) to an almost sure one. More precisely, if we look at $Z_{N,\beta_N}^{\omega,h}$ as a function of $\mathcal{P}^{(N)} := (x, \omega_x)_{x \in [-N,N]^d \cap \mathbb{Z}^d}$ and the continuous limit $\widehat{\mathcal{T}}_\beta = \widehat{\mathcal{T}}_\beta(\mathcal{P})$, then by [27, Cor. 5.12], the convergence in distribution in Theorem 2.2 implies that there exist some random elements $\widetilde{\mathcal{P}}^{(N)}$ and $\widetilde{\mathcal{P}}$ defined on the same probability space such that $\widetilde{\mathcal{P}}^{(N)} \stackrel{(d)}{=} \mathcal{P}^{(N)}$, $\widetilde{\mathcal{P}} \stackrel{(d)}{=} \mathcal{P}$ and such that $\lim_{N \rightarrow \infty} \frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\beta_N}^{\omega,h}(\widetilde{\mathcal{P}}^{(N)}) = \widehat{\mathcal{T}}_\beta(\widetilde{\mathcal{P}})$ almost surely.

It therefore makes sense to work conditionally on $\widehat{\mathcal{T}}_\beta > 0$, even at the discrete level. We have the following corollary, which says that conditionally on $\widehat{\mathcal{T}}_\beta > 0$, transversal fluctuations are of order N .

Proposition 2.4. Let $\alpha \in (0, d)$ and let $(\beta_N)_{N \geq 1}$ be such that $\lim_{N \rightarrow \infty} N^{\frac{d-\alpha}{\alpha}} \beta_N = \beta \in (0, +\infty]$. Then $(S_n)_{0 \leq n \leq N}$ has transversal fluctuations of order N under $\mathbf{P}_{N,\beta_N}^{\omega,h}(\cdot | \widehat{\mathcal{T}}_\beta > 0)$.

Let us now give some property on the variational problem $\widehat{\mathcal{T}}_\beta$, and in particular state that it is well-defined.

Proposition 2.5. When $\alpha \in (0, d)$, the variational problem $\widehat{\mathcal{T}}_\beta$ defined in (2.1) is a.s. finite for all $\beta \in (0, +\infty]$. Moreover, letting $\beta_c = \beta_c(\mathcal{P}) := \inf\{\beta > 0 : \widehat{\mathcal{T}}_\beta > 0\}$, we have that $\mathbb{P}(\beta_c \in (0, \infty)) = 1$ if $\alpha \in (0, \frac{d}{2})$ and $\mathbb{P}(\beta_c = 0) = 1$ if $\alpha \in (\frac{d}{2}, d)$.

Remark 2.6. We could have formulated the convergence in distribution in Theorem 2.2 as follows:

$$\frac{1}{N} \log Z_{N,\beta_N}^{\omega,h} \xrightarrow{(d)} \sup_{s \in \mathcal{D}, \widehat{\text{Ent}}(s) < +\infty} \{\beta \pi(s) - \widehat{\text{Ent}}(s)\} = \beta \widehat{\mathcal{T}}_\beta, \quad \text{as } N \rightarrow \infty.$$

However, in the case $\beta = +\infty$, this would only give that $\frac{1}{N} \log Z_{N,\beta_N}^{\omega,h}$ goes to $+\infty$. The formulation of Theorem 2.2 allows us to treat the case $\beta = +\infty$, in which $\beta_N N^{d/\alpha}$ is much larger than N — in particular, it includes the case $\gamma < \frac{d-\alpha}{\alpha}$. In that case, the entropy term disappears in $\widehat{\mathcal{T}}_\infty$, even though the constraint that paths have a finite entropy $\widehat{\text{Ent}}(s)$ has to remain (for instance to avoid having paths of length larger than 1).

2.2 Region B: $\alpha \in (\frac{d}{2}, d)$, $\frac{d-\alpha}{\alpha} < \gamma < \frac{d}{2\alpha}$

In region **B**, our result is the analogous to Theorem 2.5 in [7] in the context of the directed polymer model in dimension $1 + 1$: this generalizes [7] to the non-directed framework and to higher dimensions. Here, we will fix the parameter h to be equal to $\mu := \mathbb{E}[\omega_0] \in (0, \infty)$, which is well defined since $\alpha > \frac{d}{2} \geq 1$.

Theorem 2.7. Let $\alpha \in (\frac{d}{2}, d)$ and $\gamma \in (\frac{d-\alpha}{\alpha}, \frac{d}{2\alpha})$. Let $(\beta_N)_{N \geq 1}$ verify $\lim_{N \rightarrow \infty} N^\gamma \beta_N = \beta \in (0, +\infty)$. Then $(S_n)_{0 \leq n \leq N}$ has transversal fluctuations of order N^ξ under $\mathbf{P}_{N,\beta_N}^{\omega,h=\mu}$, where $\xi := \frac{\alpha(1-\gamma)}{2\alpha-d} \in (\frac{1}{2}, 1)$. Additionally, we have the following convergence in distribution, as $N \rightarrow \infty$,

$$\frac{1}{N^{2\xi-1}} \log Z_{N,\beta_N}^{\omega,h=\mu} \xrightarrow{(d)} \mathcal{T}_\beta := \sup_{s \in \mathcal{D}, \text{Ent}(s) < +\infty} \{\beta \pi(s) - \text{Ent}(s)\}, \quad (2.2)$$

where $\pi(s)$ and $\text{Ent}(s)$ are defined in (1.8) and (1.11) respectively.

Let us stress that the choice of $h = \mu$ is crucial because the contribution of the small values of the environment (that is $\beta_N \omega_x \leq 1$) to the partition function is negligible only if we center the random variables ω_x . This centering term is also needed in the directed case, cf. Section 4.2 Eq. (4.27) of [7].

The fact that the variational problem \mathcal{T}_β is well-defined in the case $\alpha \in (\frac{d}{2}, d)$ is non-trivial, and relies on (non-directed) entropic last-passage percolation (E-LPP) estimates, introduced in [9]. Properties of the variational problem \mathcal{T}_β are summarized in the

following Proposition, which is the analogous of [8, Thm. 2.4] (and generalizes it to the non-directed case and to higher dimensions).

Proposition 2.8. *If $\alpha \in (\frac{d}{2}, d)$, the variational problem \mathcal{T}_β defined in (2.2) is a.s. positive and finite for all $\beta \in (0, +\infty)$. Moreover, $\mathbb{E}[(\mathcal{T}_\beta)^\kappa] < \infty$ for any $\kappa < \alpha - d/2$, and we have the scaling relation*

$$\mathcal{T}_\beta \stackrel{(d)}{=} \beta^{\frac{2\alpha}{2\alpha-d}} \mathcal{T}_1. \quad (2.3)$$

On the other hand, if $\alpha \in (0, \frac{d}{2}]$, we have $\mathcal{T}_\beta = +\infty$ a.s., for all $\beta > 0$.

Remark 2.9. If we consider more general sequences β_N , then the corresponding transversal fluctuations r_N are given by the relation (1.5), that is $r_N \sim (N\beta_N)^{\frac{\alpha}{2\alpha-d}}$. In analogy with [7, Thms. 2.5–2.7], we should find three different scaling limits in region **B**, according to whether $\sqrt{N \log N} \ll r_N$ (corresponding to the bulk of region **B** Theorem 2.7 above), $r_N \asymp \sqrt{N \log N}$ or $\sqrt{N} \ll r_N \ll \sqrt{N \log N}$ (corresponding to boundary regions between region **B** and region **C**). We have chosen here to consider only pure powers for β_N (and for r_N) in order to keep the exposition clearer — the arguments from [7] could be adapted here but it would significantly lengthen the paper.

2.3 Region C: $\alpha \in (0, d)$, $\gamma \geq \max\{\frac{d-\alpha}{\alpha}, \frac{d}{2\alpha}\}$

In region **C**, we will show that transversal fluctuations of the polymer are of order $N^{1/2}$. As a preliminary remark, let us stress that for any $x \in \mathbb{R}^d \setminus \{0\}$ (omitting integer parts to lighten notation), we have the following asymptotics for $\mathbf{P}(x\sqrt{N} \in \mathcal{R}_N)$. Setting $v_n := \log N$ if $d = 2$ and $v_n = N^{\frac{d}{2}-1}$ if $d \geq 3$, we have

$$\lim_{N \rightarrow \infty} v_N \mathbf{P}(x\sqrt{N} \in \mathcal{R}_N) = f(x) := \begin{cases} \int_{\|x\|^2/2}^{\infty} u^{-1} e^{-u} du & \text{if } d = 2, \\ 2\lambda_d \int_0^1 \rho_d(u, x) du & \text{if } d \geq 3, \end{cases} \quad (2.4)$$

where $\lambda_d := \mathbf{P}(S_n \neq 0, \forall n \geq 1)$ is the escape probability and $\rho_d(t, x) := (2\pi t/d)^{-d/2} e^{-\frac{d}{2t}\|x\|^2}$ is the heat kernel of a d -dimensional Brownian motion with covariance matrix $\frac{1}{d}\text{Id}$. The asymptotics (2.4) can be derived from Uchiyama's results [32]: see [32, Thm. 1.6] for the case of dimension $d = 2$ and [32, Thm. 1.7] for dimensions $d \geq 3$.

In region **C**, the cases $\alpha \in (0, \frac{d}{2})$, $\gamma \geq \frac{d-\alpha}{\alpha}$ and $\alpha \in (\frac{d}{2}, d)$, $\gamma \geq \frac{d}{2\alpha}$ need to be treated separately, similarly to the case of the directed polymer model [7].

Case $\alpha \in (\frac{d}{2}, d)$

We start with the case $\alpha \in (\frac{d}{2}, d)$, $\gamma \geq \frac{d}{2\alpha}$, which is easier to state. Let us introduce some quantities that arise in the limit: we stress that different objects arise depending on the dimension.

Proposition 2.10. *In dimension $d \geq 5$, define*

$$\mathcal{X} := \sum_{x \in \mathbb{Z}^d} (\omega_x - \mu) \mathbf{P}(x \in \mathcal{R}_\infty) = \mathbf{E} \left[\sum_{x \in \mathcal{R}_\infty} (\omega_x - \mu) \right], \quad (2.5)$$

where we recall that $\mu = \mathbb{E}[\omega_0]$. Then the random variable \mathcal{X} is well-defined if $\alpha > \frac{d}{d-2}$.

Proposition 2.11. *In dimensions $d = 2, 3$ and $\alpha \in (\frac{d}{2}, 2)$, define*

$$\mathcal{W}_\beta = \int_{\mathbb{R}^d \times \mathbb{R}_+} \frac{1}{\beta} (e^{\beta w} - 1 - \beta w) f(x) \mathcal{P}(dx, dw) + \int_{\mathbb{R}^d \times \mathbb{R}_+} w f(x) (\mathcal{P} - \eta)(dx, dw), \quad (2.6)$$

with $f(x)$ defined in (2.4) and $\eta(dx, dw) = \alpha w^{-\alpha-1} dx dw$ the intensity measure for \mathcal{P} . If $\alpha \in (\frac{d}{2}, 2)$, then for any $\beta \in (0, \infty)$, the random variable \mathcal{W}_β in (2.6) is almost surely finite.

We are now ready to state our results in the case $\alpha \in (\frac{d}{2}, d)$, $\gamma \geq \frac{d}{2\alpha}$. It is the analogous of Theorem 1.4 in [19] for the directed polymer model in dimension $1 + 1$: it generalizes [19] to the non-directed case and to higher dimensions. Here again, we need to fix $h = \mu := \mathbb{E}[\mu_0]$.

Theorem 2.12. *Let $\alpha \in (\frac{d}{2}, d)$ and let $(\beta_N)_{N \geq 1}$ be such that $\lim_{N \rightarrow \infty} N^{\frac{d}{2\alpha}} \beta_N = \beta \in [0, +\infty)$. Then $(S_n)_{0 \leq n \leq N}$ has transversal fluctuations of order $N^{1/2}$ under $\mathbf{P}_{N, \beta_N}^{\omega, h=\mu}$. We also have the following convergences in distribution, as $N \rightarrow \infty$:*

• *If $\alpha \in (2 \vee \frac{d}{2}, d)$ (in particular $d \geq 3$), then setting $a_N = N^{1/4}$ for $d = 3$, $a_N = (\log N)^{1/2}$ for $d = 4$, and $a_N = 1$ for $d \geq 5$ we have:*

$$\frac{1}{a_N \beta_N} \left(\log Z_{N, \beta_N}^{\omega, h=\mu} - \frac{1}{2} \mathbb{V}\text{ar}(\omega) \beta_N^2 \mathbf{E}[|\mathcal{R}_N|] \right) \xrightarrow{(d)} \begin{cases} \mathcal{N}(0, \sigma_d^2 \mathbb{V}\text{ar}(\omega)) & \text{if } d = 3, 4, \\ \mathcal{X} & \text{if } d \geq 5, \end{cases} \quad (2.7)$$

for some explicit constant σ_d ; recall \mathcal{X} has been defined in (2.5).

• *If $\alpha \in (\frac{d}{2}, 2)$ (in particular $d = 2$ or $d = 3$), then setting $v_n := \log N$ for $d = 2$ and $v_n = N^{1/2}$ for $d = 3$, we have*

$$\frac{v_N}{\beta_N N^{\frac{d}{2\alpha}}} \log Z_{N, \beta_N}^{\omega, h=\mu} \xrightarrow{(d)} \mathcal{W}_\beta, \quad (2.8)$$

with \mathcal{W}_β defined in (2.6) — if $\beta = 0$ the first term in (2.6) is set to zero, see Proposition 2.14 below.

Remark 2.13. Let us stress that when $\alpha > d/2$, in dimension $d \geq 3$ we have $\beta_N^2 \mathbf{E}[|\mathcal{R}_N|] \sim c_\beta N^{1-d/\alpha}$ (assume $\beta > 0$ to simplify), so in particular it goes to 0 as $N \rightarrow \infty$. One can easily check that we always have $(a_N \beta_N)^{-1} \times \beta_N^2 \mathbf{E}[|\mathcal{R}_N|] \rightarrow +\infty$ when $\alpha > 2 \vee \frac{d}{2}$ (and thus, $d \geq 3$), so the centering in (2.7) is non-trivial.

Case $\alpha \in (0, \frac{d}{2})$

In this second case, recall Theorem 2.2 and Proposition 2.5. If $\alpha \in (0, \frac{d}{2})$ and $\beta_N \sim \beta N^{-(d-\alpha)/\alpha}$ as $N \rightarrow \infty$, then the scaling limit of the log-partition function has been identified as $\widehat{\mathcal{T}}_\beta$. However, when $\alpha \in (0, \frac{d}{2})$ we have $\widehat{\mathcal{T}}_\beta = 0$ when $\beta \leq \beta_c$, with $\beta_c = \beta_c(\mathcal{P}) > 0$ a.s.: in that case, the scaling limit is thus trivial. Our next result shows that when $\beta \leq \beta_c$, then transversal fluctuations are necessarily of order \sqrt{N} ; this has to be compared with Proposition 2.4 which asserts that transversal fluctuations are of order N when $\beta > \beta_c$. This shows that when $\alpha \in (0, \frac{d}{2})$ we cannot have intermediate fluctuations between $N^{1/2}$ and N : the system exhibits a sharp phase transition at the critical point $\beta = \beta_c$. Our result is the analogous of Theorem 2.12 in [7], and generalizes it to the non-directed case and to higher dimensions.

Let us define another quantity, analogous to \mathcal{W}_β defined in (2.6).

Proposition 2.14. *For $\alpha \in (0, 1) \cup (1, 2)$, define*

$$\mathcal{W}_0 = \begin{cases} \int_{\mathbb{R}^d \times \mathbb{R}_+} w f(x) (\mathcal{P} - \eta)(dx, dw), & \text{if } \alpha \in (1, 2), \\ \int_{\mathbb{R}^d \times \mathbb{R}_+} w f(x) \mathcal{P}(dx, dw), & \text{if } \alpha \in (0, 1), \end{cases} \quad (2.9)$$

with $f(x)$ defined in (2.4) and $\eta(dx, dw) = \alpha w^{-\alpha-1} dx dw$ the intensity measure for \mathcal{P} . If $\alpha \in (0, 1)$, then \mathcal{W}_0 is a.s. finite in any dimension d . If $\alpha \in (1, 2)$, then \mathcal{W}_0 is a.s. finite if and only if $\alpha < \frac{d}{d-2}$.

Recall that by an extended version of Skorokhod representation theorem [27, Cor. 5.12] (cf. Remark 2.3), it makes sense to work conditionally on $\widehat{\mathcal{T}}_\beta > 0$ or $\widehat{\mathcal{T}}_\beta = 0$, even at the discrete level. If $\alpha > 1$, we need to take $h = \mu := \mathbf{E}[\omega_0]$ and if $\alpha < 1$, then h may be any real number.

Theorem 2.15. Let $\alpha \in (0, \frac{d}{2})$, and let $(\beta_N)_{N \geq 1}$ be such that $\lim_{N \rightarrow \infty} N^{\frac{d-\alpha}{2}} \beta_N = \beta \in [0, +\infty)$. Then conditionally on the event $\{\widehat{T}_\beta = 0\}$ (i.e. $\beta \leq \beta_c$) the polymer $(S_n)_{0 \leq n \leq N}$ has transversal fluctuations of order $N^{1/2}$ under $\mathbf{P}_{N, \beta_N}^{\omega, h}$. Moreover, conditionally on $\{\widehat{T}_\beta = 0\}$, we have the following convergences in distribution, as $N \rightarrow \infty$.

- If $\alpha \in (\frac{d}{d-2}, \frac{d}{2})$ (in particular $d \geq 5$), then we have

$$\frac{1}{\beta_N} \log Z_{N, \beta_N}^{\omega, h=\mu} \xrightarrow{(d)} \mathcal{X}, \quad (2.10)$$

with \mathcal{X} defined in (2.5).

- If $\alpha < \min(\frac{d}{2}, \frac{d}{d-2})$ (in particular $\alpha < 2$) and $\alpha \neq 1$, then

$$\frac{v_N}{\beta_N N^{\frac{d}{2\alpha}}} \log Z_{N, \beta_N}^{\omega, h} \xrightarrow{(d)} \mathcal{W}_0, \quad (2.11)$$

where $v_n := \log N$ if $d = 2$ and $v_n = N^{\frac{d}{2}-1}$ if $d \geq 3$, with \mathcal{W}_0 defined in (2.9).

Remark 2.16. The case $\alpha = 1$ in (2.11) could be treated similarly to what is done in [19, Thm 1.4]. In the case $\alpha = 1$, a centering term for $\log Z_{N, \beta_N}^{\omega, h}$ is needed in (2.11) and the scaling limit should be $\mathcal{W}_0 = \int_{\mathbb{R}^d \times (0, 1]} \omega f(x) (\mathcal{P} - \eta)(dx, dw) + \int_{\mathbb{R}^d \times (1, \infty)} \omega f(x) \mathcal{P}(dx, dw)$. Since this is fairly technical, we prefer to omit the details for simplicity.

2.4 Comparison with directed polymers

As we stressed in the introduction of this paper, the model (1.1) is closely related to the directed polymer model — our results can be seen as an extension of existing results to a non-directed setting and to higher dimension. Let us now briefly discuss this relation by comparing the techniques exploited in the present article with the one developed for the directed polymer.

- In Region **A** our results extend [2, 24], where the authors considered the directed polymer in dimension $1 + 1$. In this region, only a few points give an energy contribution to the variational problem: the random walk linearly interpolates between these points to get the maximal energy reward to compensate the entropy cost. The fact that we can approximate the problem by considering only a finite number of points allows us to extend and exploit the techniques introduced for the directed case: the scheme of the proof is very close to the one in [2].

- In Region **B**, our results extend the analysis performed in [7] to the non-directed framework. In particular our techniques are based on the non-directed Entropy-controlled last passage percolation (E-LPP) introduced in [9] (as an extension of the directed E-LPP of [8]), a crucial step consisting in showing that we can restrict the partition function to trajectories staying at scale N^ξ . The extension to a non-directed setting and to higher dimension is the main novelty here, as discussed in Appendix A. Overall, the proof follows the same scheme as in [7], but several adaptations are needed, with some tedious technicalities.

- In region **C**, we perform a polynomial expansion analysis, in the same spirit as in [19] for the directed polymer model with heavy-tail disorder. The crucial difference here is that in our non-directed case the geometry of the range plays a central role in the behavior of the model (and so in the polynomial expansion analysis). We therefore have to use here a local central limit theorem for the range, considering the probability that a point at scale \sqrt{N} belongs to the range \mathcal{R}_N , i.e. $\mathbf{P}(x\sqrt{N} \in \mathcal{R}_N)$. Even if the proofs in region **C** start from the same idea as in [19], the technical treatment of the polynomial expansion requires new ideas and becomes more technical and dimension dependent. Let us stress that the case of dimension $d = 2$ requires particular care, because then

$\mathbf{P}(x\sqrt{N} \in \mathcal{R}_N)$ scales as $\log N$, compared to the polynomial behavior $N^{\frac{d}{2}-1}$ in dimension $d \geq 3$.

To conclude, let us comment on the limiting random variables that appear in Theorems 2.12-2.15; note that they indeed depend on the dimension. The random variable \mathcal{W}_β (see Propositions 2.11-2.14) appears when $\alpha < 2$, in dimensions $d = 2, 3$ or when $\beta = 0$: it is the analogous of $\mathcal{W}_\beta^{(\alpha)}$ defined in [19, p. 4011] (see also [19, Thm. 1.4]). When $\alpha > 2$, then in dimensions $d = 3, 4$, analogously to [19, Thm. 1.2] a normal random variable appears, whereas in dimension $d \geq 5$ a new random variable \mathcal{X} pops up (see Proposition 2.10 for the definition of \mathcal{X}), which to our knowledge has no analogous in the literature.

2.5 Further comments and conjectures

Our article solves completely the case $\alpha \in (0, d)$, but there are several aspects that remain to be tackled:

- In Region **B** and in Region **C** with $\alpha > \min\{2, \frac{d}{2}, \frac{d}{d-2}\}$, the parameter h is needed to be set equal to μ to center the environment: the case of a general h should be investigated;
- More precise statements on the convergence of paths could be extracted from our results, in particular one could try to prove in some cases a localization of the trajectories near an optimal path;
- The case $\alpha > d$ is still a challenge: the wandering exponent ξ is not known — except in the intermediate disorder regime (Region **C**), where $\xi = 1/2$ (but one still has to understand the correct scaling for β_N).

In this section, we develop further on these aspects, we comment their relation with the literature, and we present some open problems and conjectures.

2.5.1 About the external field h

It is not hard to see that in Region **A**, the parameter h is unimportant. On the other hand, for Region **B** and for Region **C** with $\alpha > \min\{2, \frac{d}{2}, \frac{d}{d-2}\}$, we can see from the proofs (cf. Sections 6-8) that a centering is needed, so h needs to be fixed equal to $\mu = \mathbb{E}[\omega_0]$. In order to make the proofs more transparent (which are already quite technical) and to simplify some non-central arguments, we have chosen to stick to a non-negative disorder: this is used in particular in Region **A** and **B** (they are related to the *discrete energy-entropy variational problems*, cf. Section 4). This also has the advantage to illustrate when a centering of the environment is crucial or not. Without the assumption of non-negativity of the disorder, the proofs may require some extra technical work, and one should still center the ω by the parameter h when needed (or take $\mathbb{E}[\omega_0] = 0$ and $h = 0$).

Furthermore, we can also consider a more general setting for our non-directed polymer model, by considering the following:

$$\frac{d\mathbf{P}_{N,\beta_N}^{\omega,h_N}}{d\mathbf{P}}(S) := \frac{1}{Z_{N,\beta_N}^{\omega,h_N}} \exp\left(\sum_{x \in \mathcal{R}_N} (\beta_N \omega_x - h_N)\right),$$

with $h_N = \hat{h}N^{-\zeta}$ and $\zeta \in \mathbb{R}$ (note that in this article, h_N is simply $\beta_N h$). The above model can be viewed as a *perturbed version of a random walk penalized by its range* and is more challenging. The case of the dimension $d = 1$ has been analyzed thoroughly, see [5]. In particular, we can expect that when h_N is large enough (at least compared to β_N), i.e. when the penalty by the size of the range dominates, then the polymer folds

into a ball (similarly to [4, 11, 20]). However, contrary to the homogeneous model where the center of the ball is random, the location of the ball may be completely determined by the environment, since the polymer tends to maximize the sum of ω 's seen in the ball—see [12] for the one-dimensional case. On the other hand, there should be a regime where a balance is found between the energy and the entropy (the penalization by the range is negligible—this is somehow the focus of this article), and a regime where a balance is found between the penalization by the range and the energy (the entropic cost of having an unusual range being out-weighted by the range penalty). For $\alpha \in (0, d)$, the phase diagram of the above model should be similar to the one found in the one-dimensional case, cf. [5].

2.5.2 About the geometry of the range

In this paper we consider the scaling limits of the logarithmic partition function and we describe the transversal fluctuations of the walk. Such results are the starting point to push further the analysis of the geometry of the range \mathcal{R}_N . In this setting we conjecture that the limits of the logarithmic partition function contain all the information to describe the geometry of the random walk. To be more precise, in Region **A** and **B** we conjecture that the supremum of the variational problem $\hat{\mathcal{T}}_\beta$ and \mathcal{T}_β is attained by some unique continuous path that we call \hat{s}^* and s^* respectively, by analogy with [2, 8]. We then conjecture that the typical range of the random walk is concentrated around these paths. More precisely, if we look at a path as a set, i.e. if we consider its support in \mathbb{R}^d , then we conjecture that under the hypothesis of Theorem 2.2, for any $\varepsilon > 0$,

$$\lim_{N \rightarrow \infty} \mathbf{P}_{N, \beta_N}^{\omega, h} \left(\frac{1}{N} \mathcal{R}_N \subset \mathfrak{B}_\varepsilon(\hat{s}^*) \right) \xrightarrow{\mathbb{P}} 1, \quad (2.12)$$

and under the hypothesis of Theorem 2.7, for any $\varepsilon > 0$,

$$\lim_{N \rightarrow \infty} \mathbf{P}_{N, \beta_N}^{\omega, h} \left(\frac{1}{N^\xi} \mathcal{R}_N \subset \mathfrak{B}_\varepsilon(s^*) \right) \xrightarrow{\mathbb{P}} 1, \quad (2.13)$$

where $\mathfrak{B}_\varepsilon(\gamma) = \{x \in \mathbb{R}^d : d(x, \gamma) < \varepsilon\}$ and \mathbb{P} is the coupling between the discrete and continuous disorder introduced in Remark 2.3.

In region **C** we expect that $\frac{1}{\sqrt{N}} \mathcal{R}_N$ converges to a continuous limit, which should be a random perturbation of a Brownian motion range (but a precise statement is harder to state).

2.5.3 About the intermediate disorder regime (region C)

In [25], Huang considers the intermediate disorder regime of the model (that is a regime where $\xi = 1/2$ but disorder has a non-trivial effect), in the case of a disorder with finite exponential moment, i.e. $\mathbb{E}[e^{\beta \omega_x}] < +\infty$ (or roughly speaking, $\alpha = +\infty$). More precisely, Huang [25] shows that, taking

$$\beta_N = \hat{\beta} N^{-1/4} \text{ if } d = 1, \quad \beta_N = \hat{\beta} N^{-1/2} \log N \text{ if } d = 2, \quad \beta_N = \hat{\beta} N^{-1/4} \text{ if } d = 3, \quad (2.14)$$

for some $\hat{\beta} > 0$ and taking $h_N = \beta_N^{-1} \log \mathbb{E}[e^{\beta_N \omega_x}]$, then $Z_{N, \beta_N}^{\omega, h_N}$ converges in distribution toward a random variable $\mathcal{Z}_{\hat{\beta}}$, given by an explicit Wiener chaos expansion.

Our Theorem 2.12 is the analogue of this result in the case $\frac{d}{2} < \alpha < d$ (and any dimension $d \geq 2$). In particular, (2.7)-(2.8) states that the choice $\beta_N = \hat{\beta} N^{-d/2\alpha}$ is the correct scaling in order to observe non-trivial fluctuations for $Z_{N, \beta_N}^{\omega, \mu}$. Analogously to what is done in [19] for the directed polymer, one could try to extend our Theorem 2.15

to the case $\alpha > d$, i.e. to study the intermediate disorder regime for all values of α . We leave this as an open problem, but let us comment on the expected results.

In dimension $d = 2$, we expect that the correct intermediate disorder scaling is $\beta_N = \hat{\beta}N^{-d/2\alpha} = N^{-1/\alpha}$ for all $1 < \alpha < 2$ (the scaling limit is then given by Theorem 2.12) and that one reaches the scaling $\beta_N = N^{-1/2} \log N$ given in (2.14) for all $\alpha > 2$, with then the scaling limit as in [25].

In dimension $d = 3$, we expect that the correct intermediate disorder scaling is $\beta_N = \hat{\beta}N^{-d/2\alpha} = N^{-3/2\alpha}$ for all $\frac{3}{2} < \alpha < 6$, with a scaling limit similar to Theorem 2.12, and that one reaches the scaling $\beta_N = N^{-1/4}$ given in (2.14) for all $\alpha > 6$, with then the same scaling limit as in [25]. A way to understand this comes from (2.7): loosely speaking, it says that in dimension $d = 3$, taking $\beta_N = \hat{\beta}N^{-3/2\alpha}$ we get that $Z_{N,\beta_N}^{\omega,\mu} \approx 1 + \beta_N N^{1/4} \mathcal{Z}$ for some random variable \mathcal{Z} . This approximation may remain valid for some $\alpha > d$ but it must fail when $\beta_N N^{1/4} = \hat{\beta}N^{-3/2\alpha+1/4}$ becomes of order 1 (that is when α reaches the value 6), at which point all terms in the polynomial chaos expansion of the partition function are of the same order.

In dimension $d \geq 4$, the correct intermediate disorder scaling should be $\beta_N = \hat{\beta}N^{-d/2\alpha}$ for any $\alpha > d/2$, with a scaling limit similar to Theorem 2.12. Indeed, the convergence (2.7) states that taking $\beta_N = \hat{\beta}N^{-d/2\alpha}$ we have $Z_{N,\beta_N}^{\omega,\mu} \approx 1 + \beta_N a_N \mathcal{Z}$ for some random variable \mathcal{Z} : since $a_N = (\log N)^{1/2}$ in dimension $d = 4$ and $a_N = 1$ in dimension $d \geq 5$, we have that $\beta_N a_N \rightarrow 0$ for any $\alpha < +\infty$. In that case, we should therefore have that $\lim_{n \rightarrow \infty} Z_{N,\beta_N}^{\omega,\mu} = 1$, with fluctuations of order $\beta_N a_N$. Additionally, for disorder with exponential moments (i.e. $\alpha = +\infty$), the approximation $Z_{N,\beta_N}^{\omega,\mu} \approx 1 + \beta_N a_N \mathcal{Z}$ suggests that: in dimension $d = 4$, one should take $\beta_N = \hat{\beta}/(\log N)^{1/2}$ to observe a non-trivial scaling limit for $Z_{N,\beta_N}^{\omega,\mu}$ (disorder is marginally relevant); in dimension $d \geq 5$, one should have that $Z_{N,\beta_N}^{\omega,\mu}$ has a non-trivial limit for small (but fixed) β — in other words, there is a weak disorder phase (disorder is irrelevant), in analogy with the directed polymer model in dimension $d \geq 3$.

2.5.4 About the transversal fluctuations in the case $\alpha > d$

Let us now consider the case of β fixed and let us recall the Flory argument presented in Section 1.2: for $\xi > 1/2$, the energetic gain of a polymer in a box of size N^ξ is of order $N^{d\xi/\alpha}$, while the entropic cost of going at distance N^ξ is of order $N^{2\xi-1}$. The energy-entropy balance leads to the prediction that $\xi = \frac{\alpha}{2\alpha-d}$. This should hold as long as $\xi \in (1/2, 1)$. In particular, if $\alpha < d$ we should have that $\xi = 1$: this is what is proven (in a more general framework) in the present paper, see Theorem 2.2 (Region A).

When $\alpha > d$, in comparison with the directed polymer literature (see e.g. [23]), we may conjecture that there exists some α_c such that for $\alpha \in (d, \alpha_c)$ we have $\xi = \frac{\alpha}{2\alpha-d}$ and for $\alpha > \alpha_c$ we have $\xi = \xi_c$, where $\xi_c = \xi_c(\beta) \geq 1/2$ is the fluctuation exponent obtained for a disorder with finite exponential moments. Such α_c should therefore solve the equation $\frac{\alpha}{2\alpha-d} = \xi_c$, that is $\alpha_c = \frac{d\xi_c}{2\xi_c-1}$; by convention $\alpha_c = +\infty$ when $\xi_c = 1/2$. This fact has indeed been proven in dimension $d = 1$ in [5], in the non-directed setting of the present paper: in that case we have $\xi = 1$ if $\alpha \in (0, 1)$, $\xi = \frac{\alpha}{2\alpha-1}$ if $\alpha \in (0, 2)$ and $\xi = \frac{2}{3}$ for all $\alpha > 2$.

The question in dimension $d \geq 2$ remains quite mysterious (both for non-directed and directed polymers), in particular because the fluctuation exponent ξ_c is unknown. It is expected that, at least in low dimensions, there is some critical value β_c below which $\xi_c = \frac{1}{2}$ (which would lead to $\alpha_c = +\infty$) and above which $\xi_c > 1/2$ (which would lead to $\alpha_c < +\infty$). In view of the previous subsection, one could argue that in dimension $d = 2, 3$ we have $\beta_c = 0$, but for now this is purely speculative. Let us also mention that

the existence of an upper critical dimension d_0 above which $\xi_c(d, \beta) = \frac{1}{2}$ for all β is still controversial, see e.g. [29]. In order to extend the analysis of the problem to $\alpha > d$, a first step would be to study the existence and dependence on the parameters of the conjectured exponent β_c .

2.6 Organization of the rest of the paper

We now present the organization of the paper and describe how the proofs are organized.

- In Section 3 we prove that the variational problems \mathcal{T}_β , $\widehat{\mathcal{T}}_\beta$ and the random variables \mathcal{X} and $\mathcal{W}_\beta, \mathcal{W}_0$ are well-defined (i.e. Propositions 2.5, 2.8, 2.10, 2.11 and 2.14). Along the way we discuss their main properties.

- In Section 4 we present the main auxiliary results about the discrete approximation of the energy-entropy variational problems \mathcal{T}_β and $\widehat{\mathcal{T}}_\beta$ that we use to prove Theorems 2.2–2.7. These results extend the analogue results obtained in [8, 9] to the d -dimensional and non-directed case; they are related to the Entropy-controlled Last-Passage Percolation (E-LPP). Some results on the *non-directed* version of the E-LPP are postponed to the Appendix A.

- In Section 5 we prove Theorem 2.2, i.e. region A. To prove this, we first reduce to a partition function restricted to the largest L weights $Z_{N, \beta_N}^{(L)}$, see (5.1), by showing that the contribution of all smaller weights is negligible, see Lemma 5.1. We then show that $\log Z_{N, \beta_N}^{(L)}$, properly rescaled, converges to the continuous variational problem restricted to the largest L weights, see Lemma 5.2.

- In Section 6 we prove Theorem 2.7. As a crucial first step, in Section 6.1, we show that we can restrict the partition function to trajectories staying at scale N^ξ , see Proposition 6.1. To achieve this, we need to show that the entropic cost to reach a distance $h_N \gg N^\xi$ cannot be overcompensated by an energetic gain at scale h_n : this is the purpose of Lemma 6.2. This relies once again on controlling the contribution of different ranges of weights to the partition function; the most technical part actually consists in controlling the contribution of intermediate weights, which requires a detailed analysis to match the energetic gain with the entropic cost. Once trajectories have been reduced to being at scale N^ξ , the convergence of the partition function is performed in Section 6.2, see Proposition 6.3; the strategy of the proof is similar to what is done in Section 5 for region A.

- In Section 7 we prove Theorem 2.12. The idea of the proof is based on a truncation of the environment (7.3) which allows us to prove that the main contribution comes from trajectories that stay at scale $N^{1/2}$. The convergence of the partition function defined on the truncated environment is presented in Lemma 7.2. Its proof is based on a polynomial expansion analysis (7.11) in which we show that the convergence is led by the first term. Different random limits arise depending on the tail exponent α and the dimension d . If $\alpha > 2$ the limit is Gaussian in small dimensions ($d = 3, 4$) and in high dimension ($d \geq 5$) the limit is described by the averaged (w.r.t. the random walk) sum of the environment on \mathcal{R}_∞ , see the definition (2.5) of \mathcal{X} . If $\alpha < 2$ the limit conveys the heavy-tail structure of the environment, see the definition (2.6) of \mathcal{W}_β .

- In Section 8 we prove Theorem 2.15. The most technical issue consists in showing that conditionally on $\{\widehat{\mathcal{T}}_\beta = 0\}$ the main contribution to the partition function comes from trajectories that stay at scale $N^{1/2}$, see Section 8.1; in particular, this shows that trajectories cannot have intermediate scale. Then, the strategy is similar to that used in Section 7, combining a truncation of the environment and using a polynomial expansion.

- Finally, we collect in Appendix several technical estimates. As mentioned above, Appendix A provides useful results for the non-directed E-LPP. In Appendix B, we collect

some simple random walk estimates: large and moderate deviations, probabilities that a given set is visited, intersection of ranges of independent walks. In Appendix C, we give estimates on exponential moments of a truncated version of the environment.

3 Properties of the limiting variational problems and random variables

In this section, we prove the well-posedness of the variational problems $\widehat{\mathcal{T}}_\beta$ and \mathcal{T}_β (in Section 3.2, together with some of their properties), of \mathcal{X} (in Section 3.3) and of \mathcal{W}_β and \mathcal{W}_0 (in Section 3.4).

3.1 Some notation: order statistics and truncated energy

We introduce some notation to describe the contribution of large weights to the variational problems $\widehat{\mathcal{T}}_\beta, \mathcal{T}_\beta$. For $q > 0$ we define

$$\mathcal{D}_q := \left\{ s \in \mathcal{D} : \text{Ent}(s) < \infty, \sup_{0 \leq t \leq 1} \|s(t)\| \leq q \right\}, \quad (3.1)$$

and for a realization of \mathcal{P} , we set $\mathcal{P}_q := \{(x, w) \in \mathcal{P}, \|x\| \leq q\}$. Then, the energy along a path $s \in \mathcal{D}_q$ is given by

$$\pi_q(s) := \sum_{(x, w) \in \mathcal{P}_q} w \mathbb{1}_{\{x \in s([0, 1])\}}. \quad (3.2)$$

In the ball $\Lambda_q := \{x \in \mathbb{R}^d, \|x\| \leq q\}$ we can also label the points of \mathcal{P} by using the *order statistics*:

$$(x, w)_{(x, w) \in \mathcal{P}_q} = (\mathbf{Y}_i^{(q)}, \mathbf{M}_i^{(q)})_{i \geq 1}, \quad (3.3)$$

where the distribution of $(\mathbf{Y}_i^{(q)}, \mathbf{M}_i^{(q)})_{i \geq 1}$ is given as follows:

- (i) $\mathbf{M}_i^{(q)} = (c_d q)^{d/\alpha} (\mathcal{E}_1 + \dots + \mathcal{E}_i)^{-1/\alpha}$, with $(\mathcal{E}_i)_{i \in \mathbb{N}}$ i.i.d. exponential random variables of parameter 1; this is the *sequence of weights*, $\mathbf{M}_i^{(q)}$ being the i -th largest weight in Λ_q . See for instance [18] or [24].
- (ii) $(\mathbf{Y}_i^{(q)})_{i \geq 1}$ are i.i.d. uniform random variables on Λ_q , independent of $(\mathcal{E}_i)_{i \in \mathbb{N}}$; these are the *location of the weights*;

We then define the truncated analogue of (3.2):

$$\pi_q^{(\ell)}(s) = \sum_{i=1}^{\ell} \mathbf{M}_i^{(q)} \mathbb{1}_{\{\mathbf{Y}_i^{(q)} \in s([0, 1])\}}, \quad (3.4)$$

and we also let $\pi_q^{(>\ell)}(s) = \pi_q(s) - \pi_q^{(\ell)}(s)$.

Remark 3.1. Notice that if $s \in \mathcal{D}$ verifies $\text{Ent}(s) \leq B$ with $B \geq 1$, then the length of the path s is bounded by $\sqrt{2B/d} \leq B$, and so we have $s \in \mathcal{D}_B$.

3.2 On the variational problems \mathcal{T}_β and $\widehat{\mathcal{T}}_\beta$

Recall that \mathcal{T}_β and $\widehat{\mathcal{T}}_\beta$ are defined respectively in (2.2) and (2.1). Here, we prove Propositions 2.5-2.8: We mostly focus on the first one (i.e. on \mathcal{T}_β): we treat $\widehat{\mathcal{T}}_\beta$ along the way, since the results follow from a simple comparison with \mathcal{T}_β (recall that $\widehat{\text{Ent}}(s) \geq \frac{1}{d} \text{Ent}(s)$ for all s), or with identical arguments.

Before we start, let us stress that for almost every realization of \mathcal{P} , the maps $\beta \mapsto \mathcal{T}_\beta$ and $\beta \mapsto \widehat{\mathcal{T}}_\beta$ are non-decreasing and continuous. The proof of this fact is identical to that of [8, Thm. 2.4] (see Section 4.5 in [8]), so we omit it.

Proof of the scaling relation (2.3)

Consider the Poisson Point Process \mathcal{P} on $\mathbb{R}^d \times \mathbb{R}_+$ with intensity $\eta(dx, dw) = \alpha w^{-\alpha-1} dx dw$. If we consider, for $a > 0$, the scaling transformation $\phi_a(x, w) = (ax, a^{d/\alpha}w)$, then we have that $\phi_a(\mathcal{P}) \stackrel{(d)}{=} \mathcal{P}$; notice also that $\pi^{\phi_a \star \mathcal{P}}(s) = a^{d/\alpha} \pi(s/a)$ and $\text{Ent}(as) = a^2 \text{Ent}(s)$. Therefore, since $\mathcal{D} = \frac{1}{a} \mathcal{D}$, by applying the scaling ϕ_a , we get

$$\sup_{s \in \mathcal{D}, \text{Ent}(s) < \infty} \{ \beta a^{-d/\alpha} \pi(s) - \text{Ent}(s) \} \stackrel{(d)}{=} \sup_{s \in \mathcal{D}, \text{Ent}(s) < \infty} \{ \beta \pi(s) - a^2 \text{Ent}(s) \},$$

so that $\mathcal{T}_{\beta a^{-d/\alpha}} \stackrel{(d)}{=} a^2 \mathcal{T}_{\beta a^{-2}}$. This implies (2.3). \square

Finiteness of \mathcal{T}_β (and of $\widehat{\mathcal{T}}_\beta$)

For any interval $[a, b)$ let us define

$$\mathcal{T}_\beta([a, b)) := \sup_{s: \text{Ent}(s) \in [a, b)} \{ \beta \pi(s) - \text{Ent}(s) \}. \quad (3.5)$$

In such a way, we have that

$$\mathcal{T}_\beta = \mathcal{T}_\beta([0, 1)) \vee \sup_{k \geq 0} \mathcal{T}_\beta([2^k, 2^{k+1})).$$

Using the scaling properties of \mathcal{P} and that having $\text{Ent}(s) \leq 2$ implies $s \in \mathcal{D}_2$ (see Remark 3.1), we get that

$$\mathcal{T}_\beta([2^k, 2^{k+1})) \stackrel{(d)}{=} \sup_{s \in \mathcal{D}, \text{Ent}(s) \in [1, 2)} \{ 2^{\frac{dk}{2\alpha}} \beta \pi(s) - 2^k \text{Ent}(s) \} \leq \beta 2^{\frac{dk}{2\alpha}} \sup_{s \in \mathcal{D}, \text{Ent}(s) < 2} \{ \pi(s) \} - 2^k. \quad (3.6)$$

The following lemma is the key result to prove the finiteness of \mathcal{T}_β . Its proof is similar to that of [8, Lemma 4.1] and is based on entropy-controlled last-passage percolation (E-LPP) estimates. We postpone its proof to Appendix A.

Lemma 3.2. *Let $\alpha \in (0, d)$. For any $0 < a < \alpha$ there is a constant $c = c_{a, \alpha}$ such that for any $t > 1$,*

$$\mathbb{P} \left(\sup_{s \in \mathcal{D}, \text{Ent}(s) < 2} \{ \pi(s) \} > t \right) \leq c t^{-a}.$$

In particular, it shows that for any $\alpha < d$ we have that $\sup_{\text{Ent}(s) < 2} \{ \pi(s) \}$ is a.s. finite.

Using Lemma 3.2 we can conclude the proof of the finiteness of \mathcal{T}_β when $\alpha \in (\frac{d}{2}, d)$. For any $t > 0$, using (3.6) and Lemma 3.2 we get that for any $a < \alpha$

$$\begin{aligned} \mathbb{P} \left(\mathcal{T}_\beta([2^k, 2^{k+1})) > t \right) &\leq \mathbb{P} \left(\sup_{s \in \mathcal{D}, \text{Ent}(s) \leq 2} \pi(s) > \beta^{-1} 2^{-\frac{dk}{2\alpha}} (t + 2^k) \right) \\ &\leq c \beta^a 2^{k \frac{da}{2\alpha}} (t + 2^k)^{-a}. \end{aligned} \quad (3.7)$$

Then, for any $t \geq 1$, a union bound gives

$$\mathbb{P}(\mathcal{T}_\beta > t) \leq \sum_{k=0}^{\infty} \mathbb{P}(\mathcal{T}_\beta([2^k, 2^{k+1})) > t) \leq c'_a \beta^a \left(t^{-a} \sum_{k=0}^{\log_2 t} 2^{k \frac{da}{2\alpha}} + \sum_{k > \log_2 t} 2^{-ak(1 - \frac{d}{2\alpha})} \right), \quad (3.8)$$

where we split the series at $k = \log_2 t$ and used that $t + 2^k \geq t$ for $k \leq \log_2 t$, and $t + 2^k \geq 2^k$ for $k > \log_2 t$. Since $\alpha > d/2$, we obtain that $\mathbb{P}(\mathcal{T}_\beta > t)$ is smaller than a constant (that depends on a, α, β) times

$$t^{-a} t^{\frac{da}{2\alpha}} + t^{-a(1 - \frac{d}{2\alpha})} = 2t^{-a(1 - \frac{d}{2\alpha})}.$$

In particular, for any $a < \alpha$ there is some constant $c' = c_{a,\alpha,\beta} > 0$ such that for any $t \geq 1$

$$\mathbb{P}(\mathcal{T}_\beta > t) \leq c' t^{-\frac{a}{\alpha}(\alpha - \frac{d}{2})}. \quad (3.9)$$

This concludes the proof that \mathcal{T}_β is a.s. finite when $\alpha \in (\frac{d}{2}, d)$. Also, since a can be chosen arbitrarily close to α , this proves that $\mathbb{E}[(\mathcal{T}_\beta)^\kappa] < \infty$ for any $\kappa < \alpha - d/2$. \square

Remark 3.3. Notice that, from the definition 2.1 of $\widehat{\mathcal{T}}_\beta$, we have for all $\beta \in (0, +\infty]$,

$$\widehat{\mathcal{T}}_\beta \leq \sup_{s \in \mathcal{D}, \widehat{\text{Ent}}(s) < +\infty} \{\pi(s)\} \leq \sup_{s \in \mathcal{D}, \text{Ent}(s) \leq d/2} \{\pi(s)\}. \quad (3.10)$$

For the second inequality we used that $J_d(x) = +\infty$ if $\|x\|_1 > 1$, and in particular if $\|x\|^2 > 1$: this shows that if the length of s is larger than 1 (in particular if $\text{Ent}(s) > d/2$), then $\widehat{\text{Ent}}(s) = +\infty$. Thanks to Lemma 3.2 (up to a scaling), this shows that for any $\alpha \in (0, d)$, $\widehat{\mathcal{T}}_\beta$ is a.s. finite for all $\beta > 0$.

Positivity of \mathcal{T}_β (and of $\widehat{\mathcal{T}}_\beta$)

Let us consider the random set, for any $k \in \mathbb{Z}$

$$\mathcal{G}_k := \left\{ (x, w) \in \mathcal{P}, 2^{2k-1} \leq \frac{d}{2} \|x\|^2 \leq 2^{2k}, \beta w > 2^{2k+1} \right\}. \quad (3.11)$$

If $\mathcal{G}_k \neq \emptyset$, let $(x, w) \in \mathcal{G}_k$ and s be the straight line from the origin to x , which verifies $\text{Ent}(s) = \frac{d}{2} \|x\|^2$. Therefore, on the event $\mathcal{G}_k \neq \emptyset$, we have

$$\mathcal{T}_\beta \geq \beta w - \frac{d}{2} \|x\|^2 \geq 2^{2k}. \quad (3.12)$$

Notice that $(|\mathcal{G}_k|)_{k \in \mathbb{Z}}$ are independent Poisson random variables with mean

$$\eta(|\mathcal{G}_k|) = \int_{2^k \sqrt{1/d} \leq \|x\| \leq 2^{k+1} \sqrt{2/d}} dx \int_{\beta^{-1} 2^{2k+1}}^\infty \alpha w^{-\alpha-1} dw = c_{\alpha,d,\beta} 2^{k(d-2\alpha)}.$$

Hence, we get that $\lim_{k \rightarrow -\infty} \mathbb{P}(\mathcal{G}_k \neq \emptyset) = 1$ if $\alpha > d/2$: this proves that $\mathbb{P}(\exists k, \mathcal{T}_\beta \geq 2^{2k}) = 1$ (by $\mathcal{T}_\beta < 2^{2k} \Rightarrow \mathcal{G}_k = \emptyset$ and the independence of $|\mathcal{G}_k|$), so that $\mathcal{T}_\beta > 0$ almost surely, for any $\alpha \in (\frac{d}{2}, d)$. \square

Remark 3.4. The same argument can be used to prove that $\widehat{\mathcal{T}}_\beta > 0$ almost surely for any $\alpha \in (\frac{d}{2}, \alpha)$, using that $J_d(x) \sim \frac{d}{2} \|x\|^2$ as $\|x\| \downarrow 0$, see Lemma B.1. This proves the last part of Proposition 2.5, i.e. that $\beta_c = 0$ a.s. when $\alpha \in (\frac{d}{2}, d)$.

Infiniteness of \mathcal{T}_β for $\alpha \in (0, \frac{d}{2}]$

As far as the case $\alpha \leq d/2$ is concerned, since $(|\mathcal{G}_k|)_{k \geq 1}$ are independent, the Borel-Cantelli lemma ensures that a.s. $\mathcal{G}_k \neq \emptyset$ for infinitely many $k \geq 1$. This proves that almost surely, $\mathcal{T}_\beta \geq 2^{2k}$ for infinitely many $k \geq 1$, that is $\mathcal{T}_\beta = +\infty$ a.s. when $\alpha \in (0, \frac{d}{2}]$. \square

Remark 3.5. Note that we have proven that $\widehat{\mathcal{T}}_\beta < +\infty$ a.s. for all $\beta > 0$, for any $\alpha \in (0, d)$. The reason is that in that case trajectories s with $\widehat{\text{Ent}}(s) < +\infty$ cannot exit the ball of radius 1; there is no contradiction with the fact that $\mathcal{T}_\beta = +\infty$ when $\alpha \leq d/2$.

Case $\alpha \in (0, \frac{d}{2})$, proof of $\beta_c > 0$

We now prove the part $\alpha \in (0, \frac{d}{2})$ in Proposition 2.5, i.e. that $\widehat{\mathcal{T}}_\beta = 0$ for β sufficiently small when $\alpha \in (0, \frac{d}{2})$, or equivalently $\beta_c > 0$ a.s. We proceed as in [31, Section 6].

Let \widehat{s}_β be a maximizer of $\widehat{\mathcal{T}}_\beta$, which a.s. exists (and is unique) — the proof of the existence and uniqueness of a maximizer is identical to that given in [8, Sec. 4.6]. In the following preliminary result we show that for any $\alpha \in (0, d)$, if β is small, then \widehat{s}_β has a small entropy (and is therefore confined in a small ball around the origin). We denote the length of s by $\|s\|_\infty$.

Lemma 3.6. *For any fixed $\alpha \in (0, d)$, \mathbb{P} -a.s., for any $\varepsilon > 0$ there exists $\beta_0 = \beta_0(\varepsilon) \in (0, 1)$ such that $\widehat{\text{Ent}}(\widehat{s}_\beta) < \varepsilon$ for all $\beta < \beta_0$; in particular, since $\widehat{\text{Ent}}(\widehat{s}_\beta) \geq \frac{1}{2}\|\widehat{s}_\beta\|_\infty^2$ we get that $\|\widehat{s}_\beta\|_\infty \leq \sqrt{2\varepsilon}$.*

Proof. We proceed by contradiction. Suppose that there exists $\varepsilon > 0$ and a sequence $(\beta_k)_{k \geq 1}$ with $\lim_{k \rightarrow \infty} \beta_k = 0$ such that $\widehat{\text{Ent}}(\widehat{s}_{\beta_k}) \geq \varepsilon$. Using (3.10), we have

$$0 \leq \widehat{\mathcal{T}}_{\beta_k} \leq \sup_{s \in \mathcal{D}, \widehat{\text{Ent}}(s) \leq d/2} \left\{ \pi(s) \right\} - \frac{\varepsilon}{\beta_k}. \quad (3.13)$$

By Lemma 3.2, the right hand side becomes negative as k gets large, which leads to a contradiction. \square

We can now deduce that there exists $\beta_c > 0$ such that $\widehat{\mathcal{T}}_\beta = 0$ if $\beta < \beta_c$. Let $\varepsilon > 0$ and $\beta_0 = \beta_0(\varepsilon) \in (0, 1)$ as in Lemma 3.6, that is $\widehat{\text{Ent}}(\widehat{s}_\beta) < \varepsilon$ for all $\beta < \beta_0$. Moreover define $\widehat{\varepsilon}_\beta := \widehat{\text{Ent}}(\widehat{s}_\beta) \leq \varepsilon$. Then, since $\widehat{\text{Ent}}(s) \geq \frac{1}{d}\text{Ent}(s)$, we get that if $\beta < \beta_0$

$$\widehat{\mathcal{T}}_\beta \leq \sup_{s \in \mathcal{D}, \widehat{\text{Ent}}(s) \leq \widehat{\varepsilon}_\beta} \left\{ \pi(s) \right\} - \frac{\widehat{\varepsilon}_\beta}{\beta} \leq \sup_{s \in \mathcal{D}, \text{Ent}(s) \leq d\widehat{\varepsilon}_\beta} \left\{ \pi(s) \right\} - \frac{\widehat{\varepsilon}_\beta}{\beta}. \quad (3.14)$$

For any $\varepsilon > 0$, since $\mathbb{P}(\beta_0(\varepsilon) > 0) = 1$, we can find some $\delta > 0$ such that $\mathbb{P}(\beta_0(\varepsilon) > \delta) > 1 - \varepsilon$. Then with probability large than $1 - \varepsilon$,

$$\widehat{\mathcal{T}}_\delta \leq \sup_{s \in \mathcal{D}, \text{Ent}(s) \leq d\widehat{\varepsilon}_\delta} \left\{ \pi(s) \right\} - \frac{\widehat{\varepsilon}_\delta}{\delta}. \quad (3.15)$$

Now, if $\widehat{\varepsilon}_\delta = 0$, then $\beta_c \geq \delta > 0$. Otherwise, if $\widehat{\varepsilon}_\delta > 0$, we use the following union bound for $\mathbb{P}(\widehat{\mathcal{T}}_\delta > 0)$

$$\begin{aligned} \mathbb{P}\left(\sup_{s, \text{Ent}(s) \leq d\widehat{\varepsilon}_\delta} \left\{ \pi(s) \right\} > \widehat{\varepsilon}_\delta/\delta\right) &= \sum_{k=0}^{\infty} \mathbb{P}\left(\sup_{s, \text{Ent}(s) \leq d\widehat{\varepsilon}_\delta} \left\{ \pi(s) \right\} > \widehat{\varepsilon}_\delta/\delta, \widehat{\varepsilon}_\delta \in (2^{-(k+1)}, 2^{-k}] \varepsilon\right) \\ &\leq \sum_{k=0}^{\infty} \mathbb{P}\left(\sup_{s, \text{Ent}(s) \leq d2^{-k}\varepsilon} \left\{ \pi(s) \right\} > 2^{-(k+1)}\varepsilon/\delta\right). \end{aligned} \quad (3.16)$$

Now, setting $t = 2^{-1-k(1-d/2\alpha)}\varepsilon^{1-d/2\alpha}/\delta$ and $\varepsilon_k = 2^{-k}\varepsilon$, we get that for any $k \geq 0$

$$\mathbb{P}\left(\sup_{s, \text{Ent}(s) \leq d2^{-k}\varepsilon} \left\{ \pi(s) \right\} > 2^{-(k+1)}\varepsilon/\delta\right) = \mathbb{P}\left(\sup_{s, \text{Ent}(s) \leq d\varepsilon_k} \left\{ \pi(s) \right\} > t(\varepsilon_k)^{d/(2\alpha)}\right) \leq ct^{-\alpha/2},$$

where for the last inequality we used the scaling ϕ_a above with $a = \sqrt{\varepsilon_k}$ to get that $\sup_{\text{Ent}(s) \leq d\varepsilon_k} \pi(s)$ has the same law as $\varepsilon_k^{d/(2\alpha)} \sup_{\text{Ent}(s) \leq d} \pi_1(s)$, together with Lemma 3.2. With the definition of t , we end up with

$$\mathbb{P}\left(\sup_{s, \text{Ent}(s) \leq d\widehat{\varepsilon}} \left\{ \pi(s) \right\} \geq \frac{\widehat{\varepsilon}}{\delta}\right) \leq c\delta^{\alpha/2}\varepsilon^{\frac{\alpha}{2}(\frac{d}{2\alpha}-1)} \sum_{k=0}^{\infty} 2^{k\frac{\alpha}{2}(1-\frac{d}{2\alpha})} \leq c'_\alpha \delta^{\alpha/2}\varepsilon^{d-\alpha/2}, \quad (3.17)$$

where we used that $\alpha < d/2$ to see that the sum is finite. Finally, we have

$$\mathbb{P}(\beta_c > 0) \geq 1 - \varepsilon - c'_\alpha \delta^{\alpha/2}\varepsilon^{d/4-\alpha/2}. \quad (3.18)$$

Letting first $\delta \rightarrow 0$ and then $\varepsilon \rightarrow 0$, we conclude that $\mathbb{P}(\beta_c > 0) = 1$ for $\alpha \in (0, \frac{d}{2})$. \square

3.3 Well-posedness of \mathcal{X}

We now prove Proposition 2.10. Notice that the condition $\alpha > d/(d-2)$ ensures that $\alpha > 1$ so in particular $\mu := \mathbb{E}[\omega_0]$ exists. Let us denote $\bar{\omega}_x := \omega_x - \mu$ for simplicity. We prove that $\mathcal{X} = \sum_{x \in \mathbb{Z}^d} \bar{\omega}_x \mathbf{P}(x \in \mathcal{R}_\infty)$ converges \mathbb{P} -a.s., using Kolmogorov's three-series theorem.

Before we turn to the control of the three series, let us stress that when $\alpha > d/(d-2)$ we have

$$\sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_\infty)^\alpha \leq \sum_{x \in \mathbb{Z}^d} G(x)^\alpha < +\infty, \quad (3.19)$$

where $G(x) = \sum_{n=0}^\infty \mathbf{P}(S_n = x)$ is the Green function. The fact that the sum is finite is due to the fact that $G(x) \sim c\|x\|^{2-d}$ as $x \rightarrow \infty$ in dimension $d \geq 3$ and that $\alpha > d/(d-2)$.

We now control the three series in Kolmogorov's three-series theorem.

(i) The first series is

$$\sum_{x \in \mathbb{Z}^d} \mathbb{P}(|\bar{\omega}_x| \mathbf{P}(x \in \mathcal{R}_\infty) \geq 1) \leq C \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_\infty)^\alpha < +\infty,$$

where we used Assumption (1.4) for the first inequality.

(ii) The second series is

$$\sum_{x \in \mathbb{Z}^d} \left| \mathbb{E} \left[\bar{\omega}_x \mathbf{P}(x \in \mathcal{R}_\infty) \mathbb{1}_{\{|\bar{\omega}| \mathbf{P}(x \in \mathcal{R}_\infty) \leq 1\}} \right] \right| \leq C \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_\infty)^\alpha < +\infty,$$

where we used that $\mathbb{E}[\bar{\omega}] = 0$ and $\mathbb{E}[\bar{\omega} \mathbb{1}_{\{|\omega| > u\}}] \leq cu^{1-\alpha}$ for any $u \geq 1$.

(iii) The third series is

$$\sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_\infty)^2 \text{Var}(\bar{\omega}_x \mathbb{1}_{\{|\bar{\omega}| \mathbf{P}(x \in \mathcal{R}_\infty) \leq 1\}}) \leq \begin{cases} \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_\infty)^\alpha < +\infty & \text{if } \alpha < 2, \\ \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_\infty)^2 \log \frac{1}{\mathbf{P}(x \in \mathcal{R}_\infty)} < +\infty & \text{if } \alpha \geq 2. \end{cases}$$

where we used that $\mathbb{E}[\bar{\omega}^2 \mathbb{1}_{\{|\bar{\omega}| \leq u\}}]$ is bounded by a constant times $u^{2-\alpha}$ if $\alpha < 2$, by a constant times $\log(1/u)$ if $\alpha = 2$ and by a constant if $\alpha > 2$. The fact that the second series is finite comes from the fact that $\mathbf{P}(x \in \mathcal{R}_\infty) \leq G(x)$ with $G(x) \sim c\|x\|^{2-d}$ as $x \rightarrow \infty$ and $2 > \frac{d}{d-2}$ with $d \geq 5$. This concludes the proof of Proposition 2.10. \square

3.4 Well-posedness of \mathcal{W}_β

In this part, we prove Propositions 2.11 and 2.14, starting with the second one. Let us stress that by (2.4), f is a radially decreasing function, with the following asymptotics:

$$f(x) = O(e^{-c\|x\|^2}), \quad \text{as } x \rightarrow \infty. \quad (3.20)$$

$$f(x) \sim c \begin{cases} \log(1/\|x\|), & \text{for } d = 2, \\ \|x\|^{2-d}, & \text{for } d \geq 3, \end{cases} \quad \text{as } x \rightarrow 0. \quad (3.21)$$

Recall that \mathcal{W}_0 and \mathcal{W}_β are defined in (2.9) and (2.6) respectively.

Proof of Proposition 2.14. We start with the case $\alpha \in (0, 1)$. By [27, Theorem 10.15], \mathcal{W}_0 is finite if and only if the following integral is finite:

$$\begin{aligned} \int_{\mathbb{R}^d \times \mathbb{R}_+} (wf(x) \wedge 1) w^{-(\alpha+1)} dx dw &= \int_{\mathbb{R}^d} f(x) \int_0^{1/f(x)} w^{-\alpha} dx dw + \int_{\mathbb{R}^d} \int_{1/f(x)}^\infty w^{-(1+\alpha)} dx dw \\ &= \left(\frac{1}{1-\alpha} + \frac{1}{\alpha} \right) \int_{\mathbb{R}^d} f(x)^\alpha dx, \end{aligned}$$

where we used that $\alpha \in (0, 1)$ to compute both integrals on the first line. This is always finite, thanks to the asymptotics (3.20)-(3.21) of f .

For the case $\alpha \in (1, 2)$, thanks again to [27, Theorem 10.15], \mathcal{W}_0 is finite if and only if

$$\int_{\mathbb{R}^d \times \mathbb{R}_+} (w^2 f(x)^2 \wedge w f(x)) w^{-(\alpha+1)} dx dw = \left(\frac{1}{2-\alpha} + \frac{1}{\alpha-1} \right) \int_{\mathbb{R}^d} f(x)^\alpha dx$$

is finite (the calculation is the same as above, using $\alpha \in (1, 2)$). Thanks to the asymptotics (3.20)-(3.21) of f , this is finite if and only if $\alpha < \frac{d}{d-2}$. \square

Proof of Proposition 2.11. Recall that we deal here with the case $\alpha \in (\frac{d}{2}, 2)$ with dimensions $d = 2, 3$. Recalling the definition (2.6) of \mathcal{W}_β , and using Proposition 2.14, we simply need to show that

$$\int_{\mathbb{R}^d \times \mathbb{R}_+} \frac{1}{\beta} (e^{\beta w} - 1 - \beta w) f(x) \mathcal{P}(dx, dw) < +\infty.$$

To simplify notation, we only treat the case $\beta = 1$; the case $\beta > 0$ is identical. Now, by [27, Theorem 10.15], we need to show that

$$\int_{\mathbb{R}^d \times \mathbb{R}_+} ((e^w - 1 - w) f(x) \wedge 1) w^{-(1+\alpha)} dx dw < +\infty. \quad (3.22)$$

Note that when $w \in [0, 1]$ we have $e^w - 1 - w \leq cw^2$: hence, the restriction of the above integral to $\mathbb{R}^d \times [0, 1]$ is bounded by

$$\int_{\mathbb{R}^d \times [0, 1]} (w^2 f(x) \wedge 1) w^{-(1+\alpha)} dx dw \leq \int_{\mathbb{R}^d} f(x) \int_0^1 w^{1-\alpha} dw dx = \frac{1}{2-\alpha} \int_{\mathbb{R}^d} f(x) dx,$$

which is finite thanks to the asymptotics (3.20)-(3.21) of f .

On the other hand, using that $e^w - 1 - w \leq e^w$, we also get that the restriction of the integral in (3.22) to $\mathbb{R}^d \times (1, \infty)$ is bounded by

$$\begin{aligned} & \int_{\mathbb{R}^d \times (1, \infty)} (e^w f(x) \wedge 1) w^{-(1+\alpha)} dx dw \\ & \leq C \int_1^\infty \left(\int_{\|x\|^2 > \frac{2w}{c}} e^w e^{-c\|x\|^2} dx \right) w^{-(1+\alpha)} dw + \int_1^\infty \left(\int_{\|x\|^2 \leq \frac{2w}{c}} dx \right) w^{-(1+\alpha)} dw \\ & \leq C' \int_1^\infty w^{-(1+\alpha)} dw + C'' \int_1^\infty w^{\frac{d}{2} - (\alpha+1)} dw. \end{aligned} \quad (3.23)$$

For the first inequality we used that there are constants C, c such that $f(x) \leq C e^{-c\|x\|^2}$ for all $\|x\|^2 \geq 2/c$. For the second inequality, we used that: (i) in the first term we have $e^w e^{-c\|x\|^2} \leq e^{-\frac{c}{2}\|x\|^2}$, which is integrable on \mathbb{R}^d ; (ii) in the second term we bounded the volume of the ball of radius $\sqrt{2w/c}$ by a constant times $w^{d/2}$. To conclude, we get that the first integral in (3.23) is always finite, while we use that $\alpha > d/2$ to get that the second integral is finite. \square

4 Discrete approximation of the variational problems

In this section, we present the main auxiliary results that we use to prove Theorems 2.2-2.7; they are also of independent interest. They extend analogous results obtained in [8, 9] to the d -dimensional case. We start with the definition of important quantities that are used throughout the rest of the article, then we state the convergence result.

4.1 Discrete energy-entropy variational problems

Let us introduce the discrete analogue of the variational problem in (2.2). It will appear naturally when considering the log-partition function with trajectories restricted to staying in a ball of radius $r_N \ll N$; when the scale is $r_N = N$ some other entropy functional need to be considered, analogously to (1.9) — we will treat that case afterwards.

Discrete entropy of a set of points.

For any set $\Delta = (x_1, \dots, x_k) \in (\mathbb{R}^d)^k$ of ordered distinct points (with a slight abuse of notation, we sometimes interpret Δ as a subset of \mathbb{R}^d), we define the *entropy* related to the set Δ by

$$\text{Ent}(\Delta) = \frac{d}{2} \left(\sum_{i=1}^k \|x_i - x_{i-1}\| \right)^2, \quad (4.1)$$

where $x_0 = 0$ by convention. Note that the entropy of a set Δ depends on the order of the points of Δ . Note also that if we consider s the linear interpolation of the points of Δ , we recover the continuous entropy (1.11).

We also introduce the entropy for a N -step random walk to visit Δ (see Lemma B.4):

$$\text{Ent}_N(\Delta) := \inf_{0=t_0 < t_1 < \dots < t_k \leq N} \sum_{i=1}^k \frac{d}{2} \frac{\|x_i - x_{i-1}\|^2}{t_i - t_{i-1}} = \frac{1}{N} \text{Ent}(\Delta), \quad (4.2)$$

The second identity is due to the fact that the infimum is attained for $t_i - t_{i-1} = \frac{\|x_i - x_{i-1}\|}{\sum_{i=1}^k \|x_i - x_{i-1}\|} N$.

Energy (and truncated energy) of a set of points.

For $r > 0$, let

$$\Lambda_r = \{x \in \mathbb{Z}^d : \|x\| \leq r\}$$

be the ball in \mathbb{Z}^d with radius r centered at the origin. We can write the random environment ω in Λ_r using its *ordered statistic*: we let $M_i^{(r)}$ be the i -th largest value of $(\omega_x)_{x \in \Lambda_r}$ and $Y_i^{(r)}$ its position. Then $(Y_i^{(r)})_{i=1}^{|\Lambda_r|}$ is a random permutation of Λ_r and

$$(\omega_x, x)_{x \in \Lambda_r} = (M_i^{(r)}, Y_i^{(r)})_{i=1}^{|\Lambda_r|}. \quad (4.3)$$

The energy collected by a set $\Delta = (x_1, \dots, x_k) \subset \Lambda_r$ and its contribution from the ℓ largest weights ($1 \leq \ell \leq |\Lambda_r|$) are defined by

$$\Omega_r(\Delta) := \sum_{i=1}^{|\Lambda_r|} M_i^{(r)} \mathbb{1}_{\{Y_i^{(r)} \in \Delta\}}, \quad \Omega_r^{(\ell)}(\Delta) := \sum_{i=1}^{\ell} M_i^{(r)} \mathbb{1}_{\{Y_i^{(r)} \in \Delta\}}, \quad (4.4)$$

where $\{y \in \Delta\}$ means that there is some $1 \leq i \leq k$ such that $x_i = y$, with a slight abuse of notation. Let us also set $\Omega_r^{(>\ell)}(\Delta) := \Omega_r(\Delta) - \Omega_r^{(\ell)}(\Delta)$.

Discrete variational problem.

Let us now define the discrete variational problem

$$\mathcal{T}_{N,r}^\beta := \max_{\Delta \subset \Lambda_r} \{\beta \Omega_r(\Delta) - \text{Ent}_N(\Delta)\}. \quad (4.5)$$

We also define the analogous variational problems restricted to the ℓ largest weights and beyond the ℓ -th largest weight by

$$\mathcal{T}_{N,r}^{\beta,(\ell)} := \max_{\Delta \subset \Lambda_r} \{ \beta \Omega_r^{(\ell)}(\Delta) - \text{Ent}_N(\Delta) \}, \quad (4.6)$$

$$\mathcal{T}_{N,r}^{\beta,(>\ell)} := \max_{\Delta \subset \Lambda_r} \{ \beta \Omega_r^{(>\ell)}(\Delta) - \text{Ent}_N(\Delta) \}. \quad (4.7)$$

The following result gives an explicit bound on the tail of the (discrete) variational problem. It will be useful to prove that small weights have a negligible contribution to the variational problems, uniformly in N . Its proof relies on E-LPP estimates and is very similar to [8, Prop. 2.6]; its proof is included in Appendix A.

Proposition 4.1. *The following hold true:*

- There exists a constant c , such that for any $N \geq 1$, $r \geq 1$, $\beta > 0$, $1 \leq \ell \leq |\Lambda_r|$ and $t > 1$,

$$\mathbb{P} \left(\mathcal{T}_{N,r}^{\beta,(\ell)} \geq tN \times (\beta r^{\frac{d}{\alpha}-1})^2 \right) \leq c t^{-\frac{\alpha d}{2(\alpha+d)}}. \quad (4.8)$$

- There exists a constant c , such that for any $N \geq 1$, $r \geq 1$, $\beta > 0$, $1 \leq \ell \leq |\Lambda_r|$ and $t > 1$,

$$\mathbb{P} \left(\mathcal{T}_{N,r}^{\beta,(>\ell)} \geq tN \times (\beta r^{\frac{d}{\alpha}-1} \ell^{\frac{1}{d}-\frac{1}{\alpha}})^2 \right) \leq c t^{-\frac{\alpha \ell d}{2(\alpha \ell + d)}}. \quad (4.9)$$

Adaptation for trajectories at scale N .

We also define entropy arising when considering trajectories at a scale N instead of N^ξ . For $\Delta = (x_1, \dots, x_k) \in \mathbb{R}^d$ and $N \geq 1$, let us define the N -step entropy

$$\widehat{\text{Ent}}_N(\Delta) := \inf_{0=t_0 < t_1 < \dots < t_k \leq N} \sum_{i=1}^k (t_i - t_{i-1}) J_d \left(\frac{x_i - x_{i-1}}{t_i - t_{i-1}} \right), \quad (4.10)$$

where J_d is the large deviation rate function for the random walk, see Lemma B.1. Also, we let $\widehat{\text{Ent}}(\Delta) := \widehat{\text{Ent}}_1(\Delta)$. Note that if we consider s the linear interpolation of the points of Δ , $\widehat{\text{Ent}}(\Delta)$ recovers the continuous entropy (1.11).

Let us stress that we have $\widehat{\text{Ent}}_N(\Delta) \geq \frac{1}{d} \text{Ent}_N(\Delta)$ for all $N \geq 1$, since $J_d(x) \geq \frac{d}{2} \|x\|^2$, cf. Lemma B.1. Then, we can define the analogous variational problems as in (4.5), (4.6) and (4.7), replacing the entropy $\text{Ent}_N(\Delta)$ with $\widehat{\text{Ent}}_N(\Delta)$. One difference here is that we only consider trajectories at scale N : we therefore take $r = N$ (notice also that $\widehat{\text{Ent}}_N(\Delta) = +\infty$ if Δ has one point outside Λ_N , recalling that $J_d(x) = +\infty$ if $\|x\|_1 > 1$). We define

$$\widehat{\mathcal{T}}_N^\beta := \max_{\Delta \subset \Lambda_N} \{ \beta \Omega_N(\Delta) - \widehat{\text{Ent}}_N(\Delta) \}, \quad (4.11)$$

$$\widehat{\mathcal{T}}_N^{\beta,(\ell)} := \max_{\Delta \subset \Lambda_N} \{ \beta \Omega_N^{(\ell)}(\Delta) - \widehat{\text{Ent}}_N(\Delta) \}, \quad (4.12)$$

$$\widehat{\mathcal{T}}_N^{\beta,(>\ell)} := \max_{\Delta \subset \Lambda_N} \{ \beta \Omega_N^{(>\ell)}(\Delta) - \widehat{\text{Ent}}_N(\Delta) \}. \quad (4.13)$$

Using that $\widehat{\text{Ent}}_N(\Delta) \geq \frac{1}{d} \text{Ent}_N(\Delta)$, we get that $\widehat{\mathcal{T}}_N^\beta \leq \frac{1}{d} \mathcal{T}_{N,N}^{d\beta}$, and similarly for $\widehat{\mathcal{T}}_N^{\beta,(\ell)}$ and $\widehat{\mathcal{T}}_N^{\beta,(>\ell)}$. Proposition 4.1 therefore remains valid for the variational problems (4.12)-(4.13), up to a change in the constants.

4.2 Convergence of the discrete variational problem to the continuous one

In this section, we prove that the discrete variational problems (4.5) and (4.11), when properly rescaled, converge to their continuous counterparts (2.2) and (2.1).

Recall the definition (3.1) of \mathcal{D}_q , the set of paths staying in Λ_q (the ball of radius q), and the definition (3.4) of $\pi_q^{(\ell)}(s)$, the contribution to $\pi(s)$ from the ℓ largest weights in Λ_q . We have the following convergences

Proposition 4.2. *Let $\alpha \in (\frac{d}{2}, d)$, let $q > 0$ and $\xi > 0$. Let β_N be such that $\lim_{N \rightarrow \infty} N^{d\xi/\alpha - (2\xi-1)}\beta_N = \beta \in (0, +\infty)$. Then, for any $\ell \in \mathbb{N}$ we have that*

$$\frac{1}{N^{2\xi-1}} \mathcal{T}_{N,qN^\xi}^{\beta_N,(\ell)} \xrightarrow[N \rightarrow \infty]{(d)} \mathcal{T}_{\beta,q}^{(\ell)} := \sup_{s \in \mathcal{D}_q} \{ \beta \pi_q^{(\ell)}(s) - \text{Ent}(s) \}. \quad (4.14)$$

We also have that, for any fixed $q > 0$,

$$\lim_{\ell \rightarrow +\infty} \mathcal{T}_{\beta,q}^{(\ell)} = \mathcal{T}_{\beta,q} := \sup_{s \in \mathcal{D}_q} \{ \beta \pi(s) - \text{Ent}(s) \} \quad \text{a.s.} \quad (4.15)$$

Finally, by monotonicity, we have $\lim_{q \rightarrow \infty} \mathcal{T}_{\beta,q} = \mathcal{T}_\beta$ almost surely.

We have the analogous statement for the discrete variational problem (4.12).

Proposition 4.3. *Let $\alpha \in (\frac{d}{2}, d)$. We let β_N such that $\lim_{N \rightarrow \infty} N^{\frac{d}{\alpha}-1}\beta_N = \beta \in (0, +\infty]$. Then, for any $\ell \in \mathbb{N}$ we have that*

$$\frac{1}{\beta_N N^{d/\alpha}} \widehat{\mathcal{T}}_N^{\beta_N,(\ell)} \xrightarrow[N \rightarrow \infty]{(d)} \widehat{\mathcal{T}}_\beta^{(\ell)} := \sup_{s \in \mathcal{D}_1, \widehat{\text{Ent}}(s) < +\infty} \{ \pi_1^{(\ell)}(s) - \frac{1}{\beta} \widehat{\text{Ent}}(s) \}. \quad (4.16)$$

We also have that $\lim_{\ell \rightarrow +\infty} \widehat{\mathcal{T}}_\beta^{(\ell)} = \widehat{\mathcal{T}}_\beta$ almost surely.

Notice that the restriction $s \in \mathcal{D}_1$ in (4.16) is harmless since we have $\widehat{\text{Ent}}(s) = +\infty$ for all $s \notin \mathcal{D}_1$.

Proof of Propositions 4.2 and 4.3. Let us define $r_N = qN^\xi$ for simplicity. We label the points Λ_{r_N} according to the ordered statistics $(M_i^{(r_N)}, Y_i^{(r_N)})_{i=1}^{|\Lambda_{r_N}|}$, see (4.3). In the same way, in the ball Λ_q we label the points of \mathcal{P} according to the ordered statistic $(\mathbf{M}_i^{(q)}, \mathbf{Y}_i^{(q)})_{i \geq 1}$, cf. (3.3). Then, by the Skorokhod representation theorem (cf. [18, Section 9.4], also in the same spirit of Remark 2.3), we have that for any fixed $\ell \in \mathbb{N}$,

$$\begin{aligned} \frac{1}{N^\xi} (Y_1^{(r_N)}, \dots, Y_\ell^{(r_N)}) &\xrightarrow[N \rightarrow \infty]{a.s.} (\mathbf{Y}_1^{(q)}, \dots, \mathbf{Y}_\ell^{(q)}), \\ \frac{1}{N^{\xi d/\alpha}} (M_1^{(r_N)}, \dots, M_\ell^{(r_N)}) &\xrightarrow[N \rightarrow \infty]{a.s.} (\mathbf{M}_1^{(q)}, \dots, \mathbf{M}_\ell^{(q)}). \end{aligned}$$

Then, for each $\Delta_N = (Y_{i_1}^{(r_N)}, \dots, Y_{i_k}^{(r_N)})$, thanks to the continuous mapping theorem, we get that $N^{-(2\xi-1)}\beta_N \Omega_{r_N}^{(\ell)}(\Delta_N)$ converges to $\beta \pi_q^{(\ell)}(s_\Delta)$, where s_Δ is the linear interpolation of $\Delta := (\mathbf{Y}_{i_1}^{(q)}, \dots, \mathbf{Y}_{i_k}^{(q)})$ (see the definitions (4.4) of $\Omega_{r_N}^{(\ell)}$ and (3.4) of $\pi_q^{(\ell)}(s)$). We also obtain that $N^{-(2\xi-1)}\widehat{\text{Ent}}_N(\Delta_N)$ converges to $\widehat{\text{Ent}}(s_\Delta)$. Since the maxima in (4.6) and (4.14) are over finitely many terms (in fact 2^ℓ), we get (4.14). The convergence (4.16) in Theorem 4.3 follows from exactly the same argument taking $r_N = N$ (here $q = 1$), and using that $N^{-1}\widehat{\text{Ent}}_N(\Delta_N)$ converges to $\widehat{\text{Ent}}(\Delta)$.

Finally, (4.15) simply follows from the monotonicity in ℓ , and the fact that $\mathcal{T}_{\beta,q} \leq \mathcal{T}_\beta$ which is finite almost surely. The limits $\lim_{q \rightarrow \infty} \mathcal{T}_{\beta,q} = \mathcal{T}_\beta$ and $\lim_{\ell \rightarrow \infty} \widehat{\mathcal{T}}_\beta^{(\ell)} = \widehat{\mathcal{T}}_\beta$ also follow by monotonicity. \square

5 Region A: proof of Theorem 2.2 and Proposition 2.4

In this section, we prove Theorem 2.2. First of all, notice that for any $h \in \mathbb{R}$ we have

$$e^{-N\beta_N|h|} Z_{N,\beta_N}^{\omega,h=0} \leq Z_{N,\beta_N}^{\omega,h} \leq e^{N\beta_N|h|} Z_{N,\beta_N}^{\omega,h=0}.$$

Hence, we get

$$-\frac{|h|}{N^{\frac{d-\alpha}{\alpha}}} + \frac{1}{\beta_N N^{\frac{d}{\alpha}}} \log Z_{N,\beta_N}^{\omega,h=0} \leq \frac{1}{\beta_N N^{\frac{d}{\alpha}}} \log Z_{N,\beta_N}^{\omega,h} \leq \frac{|h|}{N^{\frac{d-\alpha}{\alpha}}} + \frac{1}{\beta_N N^{\frac{d}{\alpha}}} \log Z_{N,\beta_N}^{\omega,h=0}.$$

Since $d > \alpha$, the terms $|h|N^{-(d-\alpha)/\alpha}$ go to 0 and we therefore only need to prove the result for $Z_{N,\beta_N}^{\omega,h=0}$. In the rest of this section, we will write $Z_{N,\beta_N} := Z_{N,\beta_N}^{\omega,h=0}$ for simplicity. We only treat the case where $\lim_{N \rightarrow \infty} \beta_N N^{(d-\alpha)/\alpha} = \beta \in (0, +\infty)$ for notational clarity; the case where $\beta = +\infty$ follows in a similar manner.

5.1 Convergence of the rescaled log-partition function

Let us prove Theorem 2.2. Our strategy is based on three steps: *Step 1.* most of the contribution to Z_{N,β_N} comes from some large disorder weights; *Step 2.* the log-partition function, when restricted to finitely many weights and rescaled by $\beta_N N^{d/\alpha}$, has a weak limit; *Step 3.* combine the first two steps to conclude the proof of the convergence.

Let us introduce some notation. Recall the order statistics defined in (4.3). For any $L \in \mathbb{N}$, we define, for $\rho \in \mathbb{R}_+$

$$Z_{N,\rho\beta_N}^{(L)} := \mathbf{E} \left[\exp \left(\rho \beta_N \sum_{i=1}^L M_i^{(N)} \mathbb{1}_{\{Y_i^{(N)} \in \mathcal{R}_N\}} \right) \right], \quad (5.1)$$

$$Z_{N,\rho\beta_N}^{(>L)} := \mathbf{E} \left[\exp \left(\rho \beta_N \sum_{i=L+1}^{|\Lambda_N|} M_i^{(N)} \mathbb{1}_{\{Y_i^{(N)} \in \mathcal{R}_N\}} \right) \right]. \quad (5.2)$$

Step 1. For any fixed L and $\eta \in (0, 1)$, by Hölder's inequality, we have

$$Z_{N,\beta_N} \leq \left(Z_{N,(1+\eta)\beta_N}^{(L)} \right)^{\frac{1}{1+\eta}} \left(Z_{N,(1+\eta^{-1})\beta_N}^{(>L)} \right)^{\frac{\eta}{1+\eta}},$$

so that for any fixed $L < N$, since the disorder is positive, we get

$$\log Z_{N,\beta_N}^{(L)} \leq \log Z_{N,\beta_N} \leq \frac{1}{1+\eta} \log Z_{N,(1+\eta)\beta_N}^{(L)} + \frac{\eta}{1+\eta} \log Z_{N,(1+\eta^{-1})\beta_N}^{(>L)}. \quad (5.3)$$

We now show that the last term in (5.3) can be made arbitrarily small compared with $\beta_N N^{d/\alpha}$ by choosing L large: this is the content of the following lemma.

Lemma 5.1. Assume that $\lim_{N \rightarrow \infty} \beta_N N^{\frac{d}{\alpha}-1} = \beta > 0$. For any $\varepsilon > 0$ and any $\rho > 0$, there are some $L_0 > 0$ and some N_0 such that for any $L > L_0$ and $N \geq N_0$ we have

$$\mathbb{P} \left(\frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\rho\beta_N}^{(>L)} > \varepsilon \right) \leq \varepsilon.$$

Proof. Notice that for any realization \mathcal{R}_N of the range, we have

$$\sum_{i=L+1}^{|\Lambda_N|} M_i^N \mathbb{1}_{\{Y_i^N \in \mathcal{R}_N\}} \leq \sup_{\Delta \subset \Lambda_N, \text{Ent}_N(\Delta) \leq \frac{d}{2}N} \Omega_N^{(>L)}(\Delta),$$

where we recall the notation (4.4) for $\Omega_N^{(>L)}(\Delta)$. We have used here that any set of points (x_1, \dots, x_k) visited by the random walk before time N must verify $\sum_{i=1}^k \|x_i - x_{i-1}\| \leq N$ and hence have a N -step entropy $\text{Ent}_N(\Delta) \leq \frac{d}{2}N$ (recall the definitions (4.1)-(4.2)). We therefore get that

$$\frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\rho\beta_N}^{(>L)} \leq \frac{\rho}{N^{d/\alpha}} \sup_{\Delta, \text{Ent}_N(\Delta) \leq \frac{d}{2}N} \Omega_N^{(>L)}(\Delta).$$

Now, we simply have to use Lemma A.4 (noting that $\text{Ent}_N(\Delta) = N^{-1}\text{Ent}(\Delta)$), which states that for any L and any $t > 1$

$$\mathbb{P}\left(\sup_{\Delta, \text{Ent}(\Delta) \leq \frac{d}{2}N^2} \Omega_N^{(>L)}(\Delta) \geq tL^{\frac{1}{d}-\frac{1}{\alpha}}N^{\frac{d}{\alpha}}\right) \leq t^{-\frac{\alpha Ld}{\alpha L+d}}.$$

Fixing t large enough so that the upper bound is smaller than ε (uniformly for $L \geq 1$) and then choosing L large enough so that $tL^{\frac{1}{d}-\frac{1}{\alpha}} \leq \varepsilon/\rho$, we obtain the conclusion of Lemma 5.1. \square

Step 2. Once the number of weights is fixed, we can prove the following convergence in distribution.

Proposition 5.2. *For any positive integer L and real number $\rho > 0$, we have the following convergence in distribution*

$$\frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\rho\beta_N}^{(L)} \xrightarrow{(d)} \widehat{\mathcal{T}}_{\rho\beta}^{(L)},$$

where $\widehat{\mathcal{T}}_{\rho\beta}^{(L)}$ is defined in (4.16).

Proof. Let us denote $\Upsilon_L := \{Y_1^{(N)}, \dots, Y_L^{(N)}\}$ the set of the (random) positions of the L largest weights in Λ_N . Then, we can write

$$Z_{N,\rho\beta_N}^{(L)} = \sum_{\Delta \subset \Upsilon_L} \exp\left(\rho\beta_N \Omega_{N,N}^{(L)}(\Delta)\right) \mathbf{P}(\mathcal{R}_N \cap \Upsilon_L = \Delta),$$

where we used the convention that the points of Δ are ordered, i.e. $\Delta = (x_1, \dots, x_k)$, and used some abuse of notation in writing $\Delta \subset \Upsilon_L$. Also, let us note that we use the notation $\mathcal{R}_N \cap \Upsilon_L = \Delta$ to state that the points in Δ are visited in the correct order by the random walk, and that no other point in Υ_L is visited. We then have the following lower and upper bounds:

$$\begin{aligned} Z_{N,\rho\beta_N}^{(L)} &\geq \exp\left(\sup_{\Delta \subset \Upsilon_L} \left\{\rho\beta_N \Omega_{N,N}^{(L)}(\Delta) + \log \mathbf{P}(\mathcal{R}_N \cap \Upsilon_L = \Delta)\right\}\right), \\ Z_{N,\rho\beta_N}^{(L)} &\leq 2^L L! \exp\left(\sup_{\Delta \subset \Upsilon_L} \left\{\rho\beta_N \Omega_{N,N}^{(L)}(\Delta) + \log \mathbf{P}(\Delta \subset \mathcal{R}_N)\right\}\right), \end{aligned}$$

where again we used the conventions that $\Delta = (x_1, \dots, x_k)$ are ordered points in Υ_L and that we denoted $\Delta \subset \mathcal{R}_N$ the fact that the points in Δ are visited in the correct order by the random walk. Recalling the definition (4.6) of the discrete variational problem $\widehat{\mathcal{T}}_N^{\beta,(\ell)}$, we can therefore rewrite

$$\begin{aligned} \log Z_{N,\rho\beta_N}^{(L)} - \widehat{\mathcal{T}}_N^{\rho\beta_N,(L)} &\geq \inf_{\Delta \subset \Upsilon_L} \left\{\widehat{\text{Ent}}_N(\Delta) + \log \mathbf{P}(\mathcal{R}_N \cap \Upsilon_L = \Delta)\right\}, \\ \log Z_{N,\rho\beta_N}^{(L)} - \widehat{\mathcal{T}}_N^{\rho\beta_N,(L)} &\leq \log(2^L L!) + \sup_{\Delta \subset \Upsilon_L} \left\{\widehat{\text{Ent}}_N(\Delta) + \log \mathbf{P}(\Delta \subset \mathcal{R}_N)\right\}. \end{aligned} \quad (5.4)$$

In view of the convergence in Proposition 4.3, we therefore only have to prove that the upper and lower bounds are negligible. More precisely, since there are only finitely many terms in the infimum (or in the supremum) and since $\beta_N N^{d/\alpha} \sim \beta N$ with $\beta > 0$, we only have to prove that for any distinct indices i_1, \dots, i_k , the (random) subset $\Delta = (Y_{i_1}^{(N)}, \dots, Y_{i_k}^{(N)})$ of Υ_L verifies

$$\frac{1}{N} |\widehat{\text{Ent}}_N(\Delta) + \log \mathbf{P}(\mathcal{R}_N \cap \Upsilon_L = \Delta)| \xrightarrow{\mathbb{P}} 0 \quad \text{and} \quad \frac{1}{N} |\widehat{\text{Ent}}_N(\Delta) + \log \mathbf{P}(\Delta \subset \mathcal{R}_N)| \xrightarrow{\mathbb{P}} 0. \quad (5.5)$$

Note that for any fixed L , for any $\varepsilon > 0$ we have

$$\mathbb{P}\left(\exists 1 \leq i, j \leq L, \text{ s.t. } \|Y_i^{(N)} - Y_j^{(N)}\| < \varepsilon N\right) \leq \binom{L}{2} \mathbb{P}\left(\|Y_1^{(N)} - Y_2^{(N)}\| < \varepsilon N\right) \leq C_L \varepsilon^d.$$

The points in Υ_L are therefore at distance at least εN from each other, with \mathbb{P} -probability larger than $1 - c\varepsilon^d$. Thanks to Lemma B.3 — and to the definition (4.10) of $\widehat{\text{Ent}}_N(\Delta)$ —, this shows that

$$\frac{1}{N} |\widehat{\text{Ent}}_N(\Delta) + \log \mathbf{P}(\Delta \subset \mathcal{R}_N)| \rightarrow 0 \quad (5.6)$$

with high \mathbb{P} -probability. Notice also that $\mathbf{P}(\mathcal{R}_N \cap \Upsilon_L = \Delta) = \mathbf{P}(\Delta \subset \mathcal{R}_N, \mathcal{R}_N \cap (\Upsilon_L \setminus \Delta) = \emptyset)$ and that $\mathbf{P}(\mathcal{R}_N \cap (\Upsilon_L \setminus \Delta) = \emptyset \mid \Delta \subset \mathcal{R}_N) \rightarrow 1$ with high \mathbb{P} -probability (using Lemma B.4 and that the points in Υ_L are distant by at least εN with high \mathbb{P} -probability). From (5.6), we therefore also get that

$$\frac{1}{N} |\widehat{\text{Ent}}_N(\Delta) + \log \mathbf{P}(\mathcal{R}_N \cap \Upsilon_L = \Delta)| \rightarrow 0,$$

with high \mathbb{P} -probability. This proves (5.5) and concludes the proof. \square

Step 3. Conclusion. Note that, as stressed in Proposition 4.3, we have $\lim_{L \rightarrow \infty} \widehat{\mathcal{T}}_{\rho\beta}^{(L)} = \widehat{\mathcal{T}}_{\rho\beta}$ almost surely. Due to the continuity in β (see Section 3.2) we also have that $\lim_{\rho \rightarrow 1} \widehat{\mathcal{T}}_{\rho\beta} = \widehat{\mathcal{T}}_\beta$ a.s.

The conclusion of the proof is then just a matter of combining the different steps in the correct order. For any fixed $\varepsilon > 0$, (i) we choose $\eta = \eta_\varepsilon \in (0, 1)$ sufficiently small (in (5.3)) and then $L_0 > 0$ so that for any $L \geq L_0$ both $\widehat{\mathcal{T}}_\beta^{(L)}$ and $\frac{1}{1+\eta} \widehat{\mathcal{T}}_{(1+\eta)\beta}^{(L)}$ are at distance less than ε from $\widehat{\mathcal{T}}_\beta$ with \mathbb{P} -probability larger than $1 - \varepsilon$; (ii) we fix some L large enough so that the conclusion of Lemma 5.1 holds, with $\rho = 1 + \eta^{-1}$. The conclusion then follows by applying Proposition 5.2 to both sides of the inequality (5.3) (by Skorokhod's representation theorem, cf. Remark 2.3, we can work as if the convergence in Proposition 5.2 is an almost sure convergence). \square

5.2 Transversal fluctuations: proof of Proposition 2.4

Since the transversal fluctuations are at most N , we therefore only have to prove that for any $\varepsilon > 0$ there exists some $\eta > 0$ such that for all N large enough

$$\mathbf{P}_{N,\beta_N}^{\omega,h} \left(\max_{1 \leq n \leq N} \|S_n\| \leq \eta N \right) \geq \varepsilon,$$

with large \mathbb{P} -probability conditionally on $\widehat{\mathcal{T}}_\beta > 0$.

The same proof as above can easily be adapted to show that we have the following convergence in distribution (it can be upgraded to an almost sure convergence by Skorokhod's representation theorem)

$$\frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\beta_N}^{\omega,h} \left(\max_{1 \leq n \leq N} \|S_n\| \leq \eta N \right) \xrightarrow{(d)} \widehat{\mathcal{T}}_{\beta,\eta} := \sup_{s \in \mathcal{D}_\eta, \widehat{\text{Ent}}(s) < +\infty} \left\{ \pi(s) - \frac{1}{\beta} \widehat{\text{Ent}}(s) \right\}, \quad (5.7)$$

where we set $Z_{N,\beta_N}^{\omega,h}(\mathcal{A}) = \mathbf{E}[\exp(\sum_{x \in \mathcal{R}_N} \beta_N(\omega_x - h)) \mathbb{1}_{\mathcal{A}}]$, and \mathcal{D}_η is defined in (3.1).

This shows in particular, using Skorokhod's representation theorem (see Remark 2.3), that

$$\begin{aligned} & \frac{1}{\beta_N N^{d/\alpha}} \log \mathbf{P}_{N,\beta_N}^{\omega,h} \left(\max_{1 \leq n \leq N} \|S_n\| \leq \eta N \right) \\ &= \frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\beta_N}^{\omega,h} \left(\max_{1 \leq n \leq N} \|S_n\| \leq \eta N \right) - \frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\beta_N}^{\omega,h} \xrightarrow{\text{a.s.}} \widehat{\mathcal{T}}_{\beta,\eta} - \widehat{\mathcal{T}}_\beta \end{aligned} \quad (5.8)$$

Noting that $\lim_{\eta \downarrow 0} \widehat{\mathcal{T}}_{\beta, \eta} = 0$ a.s., then conditionally on having $\widehat{\mathcal{T}}_{\beta} > 0$ we can choose η small enough so that $\widehat{\mathcal{T}}_{\beta, \eta} - \widehat{\mathcal{T}}_{\beta}$ is negative, with high (conditional) \mathbb{P} -probability. From the convergence in (5.8), on the event $\widehat{\mathcal{T}}_{\beta, \eta} - \widehat{\mathcal{T}}_{\beta} < 0$ we have that $\log \mathbf{P}_{N, \beta_N}^{\omega, h}(\max_{1 \leq n \leq N} \|S_n\| \leq \eta N)$ goes to $-\infty$, that is $\mathbf{P}_{N, \beta_N}^{\omega, h}(\max_{1 \leq n \leq N} \|S_n\| \leq \eta N)$ goes to 0 (exponentially fast). This concludes the proof. \square

6 Region B: proof of Theorem 2.7

In this section we prove Theorem 2.7. Recall that we choose $h = \mu$: we denote $\mathbf{P}_{N, \beta_N}^{\omega, h=\mu}$ and $Z_{N, \beta_N}^{\omega, h=\mu}$ by $\bar{\mathbf{P}}_{N, \beta_N}$ and \bar{Z}_{N, β_N} respectively. Recall that $\alpha \in (\frac{d}{2}, d)$ and that $\lim_{N \rightarrow \infty} \beta_N N^{-\gamma} = \beta \in (0, +\infty)$ with $\gamma \in (\frac{d-\alpha}{\alpha}, \frac{d}{2\alpha})$. Define $\xi := \frac{\alpha(1-\gamma)}{2\alpha-d}$, which turns out to be the end-to-end critical exponent, and note that we have $\xi \in (\frac{1}{2}, 1)$ for the range of parameters considered.

Let $M_N := \max_{1 \leq n \leq N} \|S_n\|_{\infty}$ (this notation is used in the rest of the paper). For any $q > 0$, we split \bar{Z}_{N, β_N} as follows

$$\bar{Z}_{N, \beta_N} = \bar{Z}_{N, \beta_N}(M_N > qN^{\xi}) + \bar{Z}_{N, \beta_N}(M_N \leq qN^{\xi}), \quad (6.1)$$

where $\bar{Z}_{N, \beta_N}(\mathcal{A}) = \mathbf{E}[\exp(\sum_{x \in \mathcal{R}_N} \beta_N(\omega_x - \mu)) \mathbb{1}_{\mathcal{A}}]$.

We divide the proof into three parts: *Step 1.* we show that $\bar{Z}_{N, \beta_N}(M_N > qN^{\xi})$ is small for q large with high \mathbb{P} -probability, which shows that $\bar{Z}_{N, \beta_N}(M_N > qN^{\xi})$ is negligible compared with \bar{Z}_{N, β_N} ; *Step 2.* we show that $\log \bar{Z}_{N, \beta_N}(M_N \leq qN^{\xi})$, when suitably rescaled, converges in distribution to $\mathcal{T}_{\beta, q}$; *Step 3.* we let $q \rightarrow \infty$ and we conclude our main result.

The main difference here with respect to Section 5 is the fact that we need to control the partition function with trajectories $M_N \geq qN^{\xi}$ (we had $\xi = 1$ in the previous section): this brings many additional technical difficulties and makes the first step much more difficult.

6.1 Step 1. Controlling $\bar{Z}_{N, \beta_N}(M_N \geq qN^{\xi})$

We prove the following estimate, slightly more general than what we need.

Proposition 6.1. *Suppose that $\beta_N N^{d\xi/\alpha} \leq cN^{2\xi-1}$ for all N , for some constant c . There exist positive constants c_1, c_2 and $\nu > 0$, such that for any sequence $A_N \geq 1$, we can find N_0 such that for any $N \geq N_0$ we have*

$$\mathbb{P}\left(\bar{Z}_{N, \beta_N}(M_N \geq A_N N^{\xi}) \geq e^{-c_1 A_N^2 N^{2\xi-1}}\right) \leq c_2 A_N^{-\nu}$$

Proof. We partition the interval $[A_N N^{\xi}, N]$ into blocks

$$B_{k, N} := [2^{k-1} N^{\xi}, 2^k N^{\xi}], \quad k = \log_2 A_N + 1, \dots, \log_2(N^{1-\xi}) \quad (6.2)$$

and we divide the partition function according to the value of M_N :

$$\bar{Z}_{N, \beta_N}(M_N \geq A_N N^{\xi}) = \sum_{k=\log_2 A_N+1}^{\log_2(N^{1-\xi})} \bar{Z}_N(M_N \in B_{k, N}). \quad (6.3)$$

By applying Cauchy–Schwarz inequality, we have

$$(\bar{Z}_{N, \beta_N}(M_N \in B_{k, N}))^2 \leq \mathbf{P}(M_N \geq 2^{k-1} N^{\xi}) \times \bar{Z}_{N, 2\beta_N}(M_N \leq 2^k N^{\xi}). \quad (6.4)$$

Then, a simple random walk estimate gives that there are some constants C_d, c_d such that for all $N \geq 1$ and all $k \geq 1$

$$\mathbf{P}(M_N \geq 2^{k-1} N^{\xi}) \leq C_d e^{-c_d 2^{2k+1} N^{2\xi-1}}. \quad (6.5)$$

To obtain this, we can for instance use a union bound on the different coordinates, to reduce to one-dimensional random walk estimates, for which such an inequality is classical, see e.g. [22].

Let $c_1 > 0$ sufficiently small such that

$$\sum_{k=\log_2 A_N+1}^{+\infty} C_d^{1/2} e^{-c_d 2^{2k-1} N^{2\xi-1}} \leq e^{-c_1 A_N^2 N^{2\xi-1}}. \quad (6.6)$$

By a union bound, we obtain

$$\begin{aligned} \mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \geq A_N N^\xi) \geq e^{-c_1 A_N^2 N^{2\xi-1}}\right) \\ \leq \sum_{k=\log_2 A_N+1}^{\log_2(N^{1-\xi})} \mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \in B_{k,N}) > C_d^{1/2} e^{-c_d 2^{2k-1} N^{2\xi-1}}\right) \\ \leq \sum_{k=\log_2 A_N+1}^{+\infty} \mathbb{P}\left(\bar{Z}_{N,2\beta_N}(M_N \leq 2^k N^\xi) > e^{c_d 2^{2k-1} N^{2\xi-1}}\right), \end{aligned} \quad (6.7)$$

where we used (6.4) and (6.5) for the last inequality. The result then follows directly from Lemma 6.2 below. \square

We formulate the following result in a general manner, since it will also be useful when $\alpha \in (0, \frac{d}{2})$. We still write $\bar{Z}_{N,\beta_N} = Z_{N,\beta_N}^{\omega, h=\mu}$, with $\mu = \mathbb{E}[\omega_0]$ when $\alpha > 1$ and μ any real number when $\alpha \in (0, 1)$.

Lemma 6.2. Assume that $\alpha \in (0, d)$ and let $(h_N)_{N \geq 1}$ be a sequence verifying $\lim_{N \rightarrow \infty} h_N^2/N = +\infty$ and $h_N \leq N$. Define $\varepsilon_N := N\beta_N h_N^{d/\alpha-2}$. For any constant c_0 , there exist constants $c > 0$ and $\nu > 0$ such that for N sufficiently large

$$\mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \leq h_N) > e^{c_0 h_N^2/N}\right) \leq c \varepsilon_N^\nu + \left(\frac{h_N^2}{N}\right)^{-\nu}. \quad (6.8)$$

As a consequence, when $\alpha > \frac{d}{2}$, if $\beta_N N^{\xi d/\alpha} \leq c N^{2\xi-1}$ then letting $h_N = q N^\xi$ we get that $\varepsilon_N \leq c q^{\frac{d}{2}-\alpha}$. Hence, one can choose $q_0 \geq 1$ such that for any $q \geq q_0$ we have

$$\mathbb{P}\left(\bar{Z}_{N,2\beta_N}(M_N \leq q N^\xi) > e^{c_0 q^2 N^{2\xi-1}}\right) \leq c q^{-\nu(2-\frac{d}{\alpha})}. \quad (6.9)$$

The proof is analogous to that of Lemma 4.1 in [7]; let us warn the reader that it is quite technical.

Proof of Lemma 6.2. Let us note that the bound is trivial if $\varepsilon_N \geq 1$. We will therefore assume that $\varepsilon_N \leq 1$ for all N ; in particular we have $\beta_N h_N^{d/\alpha} \leq h_N^2/N$.

Let us split $\bar{Z}_{N,\beta_N}(M_N \leq h_N)$ into three pieces, that we will control separately. By Hölder's inequality (using also that $\mu > 0$), we have

$$\log \bar{Z}_{N,\beta_N}(M_N \leq h_N) \leq \frac{1}{3} \log Z_{N,3\beta_N}^{(>Q)} + \frac{1}{3} \log Z_{N,3\beta_N}^{((1,Q])} + \frac{1}{3} \log \bar{Z}_{N,3\beta_N}^{(\leq 1)} \quad (6.10)$$

where we defined, for any $\rho \in \mathbb{R}_+$,

$$Z_{N,\rho\beta_N}^{(>Q)} := \mathbb{E}\left[\exp\left(\sum_{x \in \mathcal{R}_N} \rho \beta_N \omega_x \mathbb{1}_{\{\beta_N \omega_x > Q\}}\right) \mathbb{1}_{\{M_N \leq h_N\}}\right], \quad (6.11)$$

$$Z_{N,\rho\beta_N}^{(1,Q]} := \mathbb{E}\left[\exp\left(\sum_{x \in \mathcal{R}_N} \rho \beta_N \omega_x \mathbb{1}_{\{\beta_N \omega_x \in (1,Q]\}}\right) \mathbb{1}_{\{M_N \leq h_N\}}\right], \quad (6.12)$$

$$\bar{Z}_{N,\rho\beta_N}^{(\leq 1)} := \mathbb{E}\left[\exp\left(\sum_{x \in \mathcal{R}_N} \rho \beta_N (\omega_x - \mu) \mathbb{1}_{\{\beta_N \omega_x \leq 1\}}\right) \mathbb{1}_{\{M_N \leq h_N\}}\right], \quad (6.13)$$

where $Q := Q_N$ is a constant depending on N and will be chosen later (it will be of the form $(h_N^2/N)^\zeta$ for some $\zeta \in (0, 1)$). We now deal with the three terms separately.

First term: (6.11).

We want to show that by properly choosing Q of the form $(h_N^2/N)^\zeta$ for a well chosen $\zeta \in (0, 1)$, we have for N large enough

$$\mathbb{P}\left(\log Z_{N,3\beta_N}^{(>Q)} \geq c_0 \frac{h_N^2}{N}\right) \leq c\varepsilon_N^\nu + \left(\frac{h_N^2}{N}\right)^{-\nu}. \quad (6.14)$$

Let $\ell := \ell_N = (h_N^2/N)^{1-\delta}$, where δ is fixed (small enough). We consider the partition function restricted to the ℓ largest weights, as follows

$$Z_{N,3\beta_N}^{(\ell)} := \mathbf{E}\left[\exp\left(\sum_{i=1}^{\ell} 3\beta_N M_i^{(h_N)} \mathbb{1}_{\{Y_i^{(h_N)} \in \mathcal{R}_N\}}\right)\right], \quad (6.15)$$

where $(M_i^{(h_N)}, Y_i^{(h_N)})_{i \geq 0}$ is the ordered statistics introduced in (4.3).

Now, since $\beta_N h_N^{d/\alpha} \leq c h_N^2/N$, we have that

$$\mathbb{P}\left(\beta_N M_\ell^{(h_N)} > Q\right) \leq \mathbb{P}\left(M_\ell^{(h_N)} > c_\beta Q N h_N^{-2} h_N^{d/\alpha}\right) \leq \left(c \ell^{1/\alpha} N Q / h_N^2\right)^{-\alpha \ell}. \quad (6.16)$$

where we have used Lemma 5.1 in [8] for the last inequality (see (A.12) in Appendix). We now choose $Q = \ell^{-\frac{1}{\alpha}(1-\delta)^{1/2}} \frac{h_N^2}{N}$; recall that $\ell = \left(\frac{h_N^2}{N}\right)^{(1-\delta)}$. In particular, note that $Q = (h_N^2/N)^\zeta$ for some $\zeta \in (0, 1)$, and so it goes to infinity as $N \rightarrow \infty$.

Since $\ell^{1/\alpha} N Q / h_N^2 \leq \ell^{\delta/3\alpha}$ if δ has been fixed small enough, we therefore get that with probability larger than $1 - (c \ell^{\delta/3\alpha})^{-\alpha \ell}$ we have

$$\{x \in \mathcal{R}_N : \beta_N \omega_x > Q\} \subset \Upsilon_\ell^{(h_N)} = \{Y_1^{(h_N)}, \dots, Y_\ell^{(h_N)}\}. \quad (6.17)$$

We stress that on this event we have $Z_{N,3\beta_N}^{(>Q)} \leq Z_{N,3\beta_N}^{(\ell)}$, and we now turn to $Z_{N,3\beta_N}^{(\ell)}$. Recall the notation (4.4) of $\Omega^\ell(\Delta)$. By Lemma B.4, we have that

$$Z_{N,3\beta_N}^{(\ell)} = \sum_{\Delta \subset \Upsilon_\ell^{(h_N)}} e^{3\beta_N \Omega_{h_N}^{(\ell)}(\Delta)} \mathbf{P}(\Delta \subset \mathcal{R}_N) \leq \sum_{\Delta \subset \Upsilon_\ell^{(h_N)}} (C_1)^{|\Delta|} \exp\left(6\beta_N \Omega_{h_N}^{(\ell)}(\Delta) - C_2 \text{Ent}_N(\Delta)\right)$$

(recall that we view $\Delta = (x_1, \dots, x_k)$ as an ordered subset of $\Upsilon_\ell^{(h_N)}$). Therefore, recalling the definition (4.6) of the variational problem $\mathcal{T}_{N,r}^{\beta,(\ell)}$, we have, setting $C_3 := 6/C_2$

$$Z_{N,3\beta_N}^{(\ell)} \leq \ell! (2C_1)^\ell \exp\left(C_2 \mathcal{T}_{N,h_N}^{C_3\beta_N,(\ell)}\right). \quad (6.18)$$

Then for N large enough, using the definition of $\ell = (h_N^2/N)^{1-\delta}$, we get

$$\log Z_{N,3\beta_N}^{(\ell)} \leq \ell(\log \ell + \log(2C_1)) + C_2 \mathcal{T}_{N,h_N}^{C_3\beta_N,(\ell)} \leq \frac{c_0}{2} \frac{h_N^2}{N} + C_2 \mathcal{T}_{N,h_N}^{C_3\beta_N,(\ell)}, \quad (6.19)$$

where c_0 is a fixed constant (appearing in (6.14)). Recalling the definition $\varepsilon_N := N\beta_N h_N^{d/\alpha-2}$, we therefore get that

$$\begin{aligned} \mathbb{P}\left(\log Z_{N,3\beta_N}^{(\ell)} \geq c_0 \frac{h_N^2}{N}\right) &\leq \mathbb{P}\left(\mathcal{T}_{N,h_N}^{C_3\beta_N,(\ell)} \geq \frac{c_0}{2C_2} \frac{h_N^2}{N}\right) \\ &\leq \mathbb{P}\left(\mathcal{T}_{N,h_N}^{C_3\beta_N,(\ell)} \geq \frac{c_0}{2C_2 C_3^2} \varepsilon_N^{-2} N \left(C_3 \beta_N (h_N)^{\frac{d}{\alpha}-1}\right)^2\right). \end{aligned} \quad (6.20)$$

Then, one simply needs to use Proposition 4.1 to get that the last term is bounded by a constant times $\varepsilon_N^{\alpha d/(\alpha+d)}$. This, together with (6.16), establishes (6.14).

Second term: (6.12).

We now show that there exists some $\zeta \in (0, 1)$, such that for any $C > 0$ and N large enough,

$$\mathbb{P} \left(\log Z_{N,3\beta_N}^{((1,Q])} \geq C \left(\frac{h_N^2}{N} \right)^\zeta \right) \leq c \left(\frac{h_N^2}{N} \right)^{-\nu}. \quad (6.21)$$

We again decompose $Z_{N,3\beta_N}^{((1,Q])}$, according to the contribution of different ranges of values in the environment. We fix δ (appearing in the definition of Q and ℓ above) small enough so that $\theta := (1 - \delta) \frac{d}{\alpha} > 1$ (recall $\alpha < d$). Then, for $j \geq 0$, we set

$$\ell_j := \left(\frac{h_N^2}{N} \right)^{\theta^j(1-\delta)} = (\ell)^{\theta^j}, \quad (6.22)$$

$$Q_j := \left(\frac{h_N^2}{N} \right) (\ell_j)^{-\frac{1}{\alpha}(1-\delta)^{1/2}} = \left(\frac{h_N^2}{N} \right)^{1 - \frac{\theta^j}{\alpha}(1-\delta)^{3/2}}, \quad (6.23)$$

where we used the definition of ℓ above. Note that $\ell_0 = \ell$, $Q_0 = Q$ and that each pair (ℓ_j, Q_j) has a similar form to (ℓ, Q) , and let us stress that $\ell_{j-1} < \ell_j$ and $Q_{j-1} > Q_j$ for any $j \geq 1$. Now, let κ be the smallest integer such that $\theta^\kappa(1 - \delta)^{3/2} \geq \alpha$, and note that for such κ we have $Q_\kappa \leq 1$.

Then, thanks to Hölder's inequality, we have

$$\log Z_{N,3\beta_N}^{((1,Q])} \leq \frac{1}{\kappa} \sum_{j=1}^{\kappa} \log Z_{N,3\kappa\beta_N}^{(j)}, \quad (6.24)$$

where we have set

$$Z_{N,3\kappa\beta_N}^{(j)} := \mathbf{E} \left[\exp \left(\sum_{x \in \mathcal{R}_N} 3\kappa\beta_N \omega_x \mathbb{1}_{\{\beta_N \omega_x \in (Q_j, Q_{j-1}]\}} \right) \mathbb{1}_{\{M_N \leq h_N\}} \right]. \quad (6.25)$$

Thanks to an union bound we only need to show that for any $j = 1, \dots, \kappa$,

$$\mathbb{P} \left(\log Z_{N,3\kappa\beta_N}^{(j)} \geq C \left(\frac{h_N^2}{N} \right)^\zeta \right) \leq c \left(\frac{h_N^2}{N} \right)^{-\nu}. \quad (6.26)$$

to get (6.21).

By the same argument used for $Z_{N,3\beta_N}^{(>Q)}$, see (6.16)-(6.17), we get that

$$\{x \in \mathcal{R}_N : \beta_N \omega_x > Q_j\} \subset \Upsilon_{\ell_j} := \{Y_1^{(h_N)}, \dots, Y_{\ell_j}^{(h_N)}\}. \quad (6.27)$$

with probability larger than $1 - (c\ell_j^{\delta/3\alpha})^{-\alpha\ell_j}$. Hence, as above, we are reduced to controlling

$$Z_{N,3\kappa\beta_N}^{(j)} \leq \mathbf{E} \left[\exp \left(3\kappa Q_{j-1} \sum_{k=1}^{\ell_j} \mathbb{1}_{\{Y_k^{(h_N)} \in \mathcal{R}_N\}} \right) \right] = \sum_{k=0}^{\ell_j} \sum_{\Delta \subset \Upsilon_{\ell_j}, |\Delta|=k} e^{3\kappa Q_{j-1}k} \mathbf{P}(\mathcal{R}_N \cap \Upsilon_{\ell_j} = \Delta). \quad (6.28)$$

We split the sum in (6.28) at some level $K = K_j := (\ell_j)^{1/d+\delta/6d}$ (this is needed in (6.30) below) and we use Lemma B.4 to bound the probability $\mathbf{P}(\Delta \subset \mathcal{R}_N) \geq \mathbf{P}(\mathcal{R}_N \cap \Upsilon_{\ell_j} = \Delta)$ for $k \geq K$. We get

$$\begin{aligned} Z_{N,3\kappa\beta_N}^{(j)} &\leq e^{3\kappa Q_{j-1}K} + \sum_{k=K}^{\ell_j} \sum_{\Delta \subset \Upsilon_{\ell_j}, |\Delta|=k} (C_1)^k \exp \left(3\kappa Q_{j-1}k - C_2 \text{Ent}_N(\Delta) \right) \\ &\leq e^{3\kappa Q_{j-1}K} + \mathcal{H}_j, \end{aligned} \quad (6.29)$$

where we have set

$$\mathcal{H}_j := \sum_{k=K}^{\ell_j} k! \binom{\ell_j}{k} (C_1)^k \exp \left(6\kappa Q_{j-1} k - C_2 \inf_{\Delta \subset \Upsilon_{\ell_j}, |\Delta| \geq k} \text{Ent}_N(\Delta) \right).$$

We now use Entropic LPP estimates to control the last infimum. We let

$$B_k := \frac{12\kappa}{C_2} Q_{j-1} k = \frac{12\kappa}{C_2} \frac{h_N^2}{N} \ell_j^{-\frac{1}{d}(1-\delta)^{-1/2}} k,$$

where we used the definition (6.23) and the fact that $\ell_{j-1} = \ell_j^{1/\theta}$ together with the definition $\theta = (1 - \delta)^{\frac{d}{\alpha}}$. Then, thanks to Theorem A.1 in Appendix, we obtain

$$\mathbb{P} \left(\inf_{\Delta \subset \Upsilon_{\ell_j}, |\Delta| \geq k} \text{Ent}_N(\Delta) \leq B_k \right) = \mathbb{P} \left(L_{\ell_j}^{(NB_k)}(h_N) \geq k \right) \leq \left(C' \frac{(NB_k)^{1/2} \ell_j^{1/d}}{h_N k} \right)^{dk},$$

so that using the definition of B_k , we end up with

$$\mathbb{P} \left(\inf_{\Delta \subset \Upsilon_{\ell_j}, |\Delta| \geq k} \text{Ent}_N(\Delta) \leq B_k \right) \leq \left(c(\ell_j)^{\frac{1}{d} - \frac{1}{2d(1-\delta)^{1/2}}} k^{-1/2} \right)^{dk} \leq (c\ell_j)^{-k\delta/4},$$

where we have used in the last inequality that $k \geq K := (\ell_j)^{1/d+\delta/6d}$ (and took δ small enough). By a union bound, this leads to

$$\mathbb{P} \left(\inf_{\Delta \subset \Upsilon_{\ell_j}, |\Delta| \geq k} \text{Ent}(\Delta) > B_k, \forall K \leq k \leq \ell_j \right) \geq 1 - \sum_{k=K}^{\infty} (c\ell_j)^{-k\delta/4} \geq 1 - (c\ell_j)^{-c_\delta \ell_j^{1/d}}. \quad (6.30)$$

Moreover, on this event, we get that for N large enough

$$\mathcal{H}_j \leq \sum_{k=K}^{\infty} k! \binom{\ell_j}{k} (C_1)^k e^{-6\kappa Q_{j-1} k} \leq \sum_{k=1}^{\infty} \ell_j^k \exp \left(-4\kappa k (h_N^2/N)^\vartheta \right),$$

where we used in the last inequality that $Q_{j-1} \geq Q_{\kappa-1} = (h_N^2/N)^\vartheta$ for some $\vartheta = \vartheta_\kappa > 0$ (recall (6.23)), and that $(h_N^2/N)^\vartheta$ goes to infinity. Now using that $\ell_j \leq \ell_\kappa \leq (h_N^2/N)^{\theta_\kappa(1-\delta)}$, we get that for N large enough

$$\mathcal{H}_j \leq \sum_{k=1}^{\infty} \exp \left(-2\kappa k (h_N^2/N)^\vartheta \right) \leq \exp \left(-\kappa (h_N^2/N)^\vartheta \right). \quad (6.31)$$

Note also that $KQ_{j-1} \leq (h_N^2/N) \ell_j^{-\delta/4d} \leq (h_N^2/N) \ell^{-\delta/4d}$ goes to 0 as $N \rightarrow \infty$: going back to (6.29), we therefore get that on the event considered in (6.30), for N large enough we have $Z_{N,3\kappa\beta_N}^{(j)} \leq c(h_N^2/N)^\zeta$.

Combining this with (6.27) and (6.30), we therefore obtain that

$$\mathbb{P} \left(Z_{N,3\kappa\beta_N}^{(j)} \geq 2 \right) \leq (c\ell_j)^{-\delta\ell_j/3} + (c\ell_j)^{-c_\delta \ell_j^{1/d}}, \quad (6.32)$$

which proves (6.26), recalling that $\ell_j \geq \ell = (h_N^2/N)^{1-\delta}$.

Third term: (6.13)

We now show that there exists a constant C such that

$$\mathbb{P} \left(\log \bar{Z}_{N,3\kappa\beta_N}^{(\leq 1)} > c_0 \frac{h_N^2}{N} \right) < C \left(\frac{h_N^2}{N} \right)^{-\nu}. \quad (6.33)$$

To bound (6.33) we note that there is some $C > 0$ such that $e^x \leq 1 + x + Cx^2$ for $|x| \leq 4$. Therefore, for N large enough (so that $3\beta_N\mu \leq 1$), we obtain

$$\bar{Z}_{N,3\beta_N}^{(\leq 1)} \leq \mathbb{E} \left[\prod_{x \in \mathcal{R}_N} (1 + 3\beta_N(\omega - \mu) \mathbb{1}_{\{\beta_N\omega_x \leq 1\}} + C'\beta_N^2(\omega_x - \mu)^2 \mathbb{1}_{\{\beta_N\omega_x \leq 1\}}) \right]. \quad (6.34)$$

Note that $\mathbb{E}[(\omega - \mu) \mathbb{1}_{\{\beta_N\omega \leq 1\}}] \leq 0$ as soon as $\beta_N^{-1} \geq \mu$ (our assumption implies that $\beta_N \rightarrow 0$). Hence, using also that $|\mathcal{R}_N| \leq N$, we get that

$$\mathbb{E}[\bar{Z}_{N,3\beta_N}^{(\leq 1)}] \leq \left(1 + C'\beta_N^2 \mathbb{E}[(\omega - \mu)^2 \mathbb{1}_{\{\beta_N\omega \leq 1\}}] \right)^N \leq \begin{cases} \exp(CN\beta_N^\alpha) & \text{if } \alpha < 2, \\ \exp(CN\beta_N^2 \log(1/\beta_N)) & \text{if } \alpha \geq 2, \end{cases} \quad (6.35)$$

where we used that the truncated expectation is bounded by a constant times $(1/\beta_N)^{2-\alpha}$ if $\alpha < 2$, by a constant times $\log(1/\beta_N)$ if $\alpha = 2$, and by a constant if $\alpha > 2$.

Then by Markov's inequality and (6.35), we have that

$$\mathbb{P} \left(\log \bar{Z}_{N,3\beta_N}^{(\leq 1)} > c_0 \frac{h_N^2}{N} \right) \leq e^{-c_0 \frac{h_N^2}{N}} \mathbb{E}[\bar{Z}_{N,3\beta_N}^{(\leq 1)}] \leq e^{-c_0 \frac{h_N^2}{N}} \times \begin{cases} \exp(CN\beta_N^\alpha), & \text{if } \alpha < 2, \\ \exp(CN\beta_N^2 \log(1/\beta_N)), & \text{if } \alpha \geq 2. \end{cases} \quad (6.36)$$

It suffices to show that $h_N^{-2} N^2 \beta_N^{2 \wedge \alpha} \log(1/\beta_N) \mathbb{1}_{\{\alpha \geq 2\}} \xrightarrow{N \rightarrow \infty} 0$. Since $\beta_N h_N^{d/\alpha} \leq ch_N^2/N$, we get that:

* If $\alpha < 2$ then

$$\frac{N^2}{h_N^2} \beta_N^\alpha \leq c' \left(\frac{h_N^2}{N} \right)^{\alpha-2} h_N^{2-d} \leq c' \left(\frac{h_N^2}{N} \right)^{\alpha-2}, \quad (6.37)$$

because $d \geq 2$.

* If on the other hand $\alpha \geq 2$, we can choose $\varepsilon > 0$ sufficiently small, such that

$$\frac{N^2}{h_N^2} \beta_N^2 \log(1/\beta_N) \leq c \frac{N^2}{h_N^2} \beta_N^{2-\varepsilon} \leq ch_N^{2(1-\frac{d}{\alpha})+\frac{d}{\alpha}\varepsilon} \left(\frac{h_N^2}{N} \right)^{-\varepsilon} \leq N^{1-\frac{d}{\alpha}+\frac{d}{2\alpha}\varepsilon} \leq \left(\frac{h_N^2}{N} \right)^{1-\frac{d}{\alpha}+\frac{d}{2\alpha}\varepsilon} \quad (6.38)$$

using that $h_N \gg N^{1/2}$ for the third inequality and that $h_N^2/N \leq N$ for the last inequality. All together, this proves (6.33).

The conclusion of Lemma 6.2 then simply follows by combining (6.14), (6.21) and (6.33). \square

6.2 Step 2: convergence of $\log \bar{Z}_N(M_N \leq qN^\xi)$

We now prove the following convergence.

Proposition 6.3. Assume that $\lim_{N \rightarrow \infty} \beta_N N^{d\xi/\alpha} N^{-(2\xi-1)} = \beta \in (0, +\infty)$. Then for all $q \in (0, +\infty)$ we have the following convergence in distribution

$$\frac{1}{N^{2\xi-1}} \log \bar{Z}_{N,\beta_N}(M_N \leq qN^\xi) \xrightarrow{(d)} \mathcal{T}_{\beta,q},$$

where $\mathcal{T}_{\beta,q}$ is defined in (4.15).

Once trajectories are restricted to staying in a ball of radius qN^ξ , the proof of the convergence is very similar to the proof in Section 5. We follow the same steps: (i) first we show that most of the contribution to the partition function comes from large disorder weights; (ii) we prove that the log-partition function, when restricted to finitely many weights, converges in distribution; (iii) we send the number of weights to infinity.

Step 2.(i)-a. First of all, we show that we can restrict the partition function to weights $\omega_x \geq 1/\beta_N$. Recall the notations $\bar{Z}_{N,\rho\beta}^{(\leq 1)}$ in (6.13), and define analogously $\bar{Z}_{N,\rho\beta}^{(>1)}$ by

replacing $\mathbb{1}_{\{\beta_N \omega_x \leq 1\}}$ with $\mathbb{1}_{\{\beta_N \omega_x > 1\}}$ inside (6.13); take $h_N = qN^\xi$. For any $\eta \in (0, 1)$, by Hölder's inequality, we have

$$\bar{Z}_{N, \beta_N}(M_N \leq qN^\xi) \leq \left(\bar{Z}_{N, (1+\eta)\beta_N}^{(>1)} \right)^{\frac{1}{1+\eta}} \left(\bar{Z}_{N, (1+\eta^{-1})\beta_N}^{(\leq 1)} \right)^{\frac{\eta}{1+\eta}}, \quad (6.39)$$

and by applying Hölder's inequality to $\bar{Z}_{N, (1-\eta)\beta_N}^{(>1)}$, we also have for any $\eta \in (0, 1)$

$$\bar{Z}_{N, \beta_N}(M_N \leq qN^\xi) \geq \left(\bar{Z}_{N, (1-\eta)\beta_N}^{(>1)} \right)^{\frac{1}{1-\eta}} \left(\bar{Z}_{N, (1-\eta^{-1})\beta_N}^{(\leq 1)} \right)^{\frac{\eta}{1-\eta}}. \quad (6.40)$$

Note that if $\beta_N \omega > 1$ and N large enough, we have

$$(1 - 2\eta)\beta_N \omega < (1 - \eta)\beta_N(\omega - \mu) \quad \text{and} \quad (1 + \eta)\beta_N(\omega - \mu) \leq (1 + \eta)\beta_N \omega,$$

so that we can replace $\bar{Z}_{N, (1+\eta)\beta_N}^{(>1)}$ and $\bar{Z}_{N, (1-\eta)\beta_N}^{(>1)}$ by $Z_{N, (1+\eta)\beta_N}^{(>1)}$ and $Z_{N, (1-2\eta)\beta_N}^{(>1)}$ respectively, where

$$Z_{N, \rho\beta_N}^{(>1)} := \mathbf{E} \left[\exp \left(\rho\beta_N \sum_{x \in \mathcal{R}_N} \omega_x \mathbb{1}_{\{\beta_N \omega_x > 1\}} \right) \mathbb{1}_{\{M_N \leq qN^\xi\}} \right]. \quad (6.41)$$

We need to control the upper bound for (6.39) and the lower bound for (6.40) (note that $1 - \eta^{-1}, \eta - 1 < 0$). Now notice that adapting (6.36), we easily get that for any $\rho > 0$, any $\varepsilon > 0$

$$\mathbb{P} \left(N^{-(2\xi-1)} \log \bar{Z}_{N, \rho\beta_N}^{(\leq 1)} > \varepsilon \right) \leq \exp(C_\rho N \beta_N^{2\wedge \alpha} \log(1/\beta_N) - \varepsilon N^{2\xi-1}),$$

which goes to 0 as $N \rightarrow \infty$ (the proof of (6.35) is carried out for $\rho = 3$ but it is easily adapted to any $\rho > 0$). For $\rho < 0$, a slight modification in (6.34)-(6.35) is needed. Note that we have $\mathbf{E}[-\beta_N(\omega - \mu) \mathbb{1}_{\{\beta_N \omega_x \leq 1\}}] = \mathbf{E}[\beta_N(\omega - \mu) \mathbb{1}_{\{\beta_N \omega_x > 1\}}] \leq C\beta_N^\alpha$, which gives an extra term in (6.35) for $\rho < 0$ and it follows that

$$\mathbb{P} \left(-N^{-(2\xi-1)} \log \bar{Z}_{N, \rho\beta_N}^{(\leq 1)} < -\varepsilon \right) \leq \exp(C_\rho N \beta_N^{2\wedge \alpha} \log(1/\beta_N) - \varepsilon N^{2\xi-1}) \xrightarrow{N \rightarrow \infty} 0.$$

From (6.39)-(6.40) we simply need to prove the convergence in distribution of $N^{-(2\xi-1)} \log Z_{N, \rho\beta_N}^{(>1)}$ to $\mathcal{T}_{\rho\beta, q}$.

Notice also that thanks to (6.21), we also have that for any $\rho > 0$

$$\frac{1}{N^{2\xi-1}} \log Z_{N, \rho\beta_N}^{(1, Q]} \longrightarrow 0, \quad \text{in probability.}$$

Using Hölder's inequality as in (6.39) and (6.40), we are therefore reduced to showing that for any $\rho > 0$ we have the following convergence in distribution

$$\frac{1}{N^{2\xi-1}} \log Z_{N, \rho\beta_N}^{(>Q)} \xrightarrow{(d)} \mathcal{T}_{\rho\beta, q}. \quad (6.42)$$

Step 2.(i)-b. We now show that we can restrict the partition function to a finite number of weights. Analogously to Section 5 (see (5.1)-(5.2)), we define for any $L \in \mathbb{N}$,

$$Z_{N, \rho\beta_N}^{(L)} := \mathbf{E} \left[\exp \left(\rho\beta_N \sum_{i=1}^L M_i^{(qN^\xi)} \mathbb{1}_{\{Y_i^{(qN^\xi)} \in \mathcal{R}_N\}} \right) \right], \quad (6.43)$$

$$Z_{N, \rho\beta_N}^{(L, \ell)} := \mathbf{E} \left[\exp \left(\rho\beta_N \sum_{i=L+1}^{\ell} M_i^{(qN^\xi)} \mathbb{1}_{\{Y_n^{(qN^\xi)} \in \mathcal{R}_N\}} \right) \right], \quad (6.44)$$

where we consider the order statistics $(M_i^{(qN^\xi)}, Y_i^{(qN^\xi)})$ in the (discrete) ball of radius qN^ξ . Recall that $\ell := (h_N^2/N)^{1-\delta} = (q^2 N^{2\xi-1})^{1-\delta}$.

Thanks to Hölder's inequality, we get that for any $\eta \in (0, 1)$, analogously to (5.3), then for N large enough we have

$$\log Z_{N,\rho\beta_N}^{(L)} \leq \log Z_{N,\rho\beta_N}^{(>Q)} \leq \frac{1}{1+\eta} \log Z_{N,(1+\eta)\beta_N}^{(L)} + \frac{\eta}{1+\eta} \log Z_{N,(1+\eta^{-1})\beta_N}^{(L,\ell)} \quad (6.45)$$

with probability going to 1 as $N \rightarrow \infty$ (more precisely on the event $\{\beta_N M_L^{(qN^\xi)} \geq Q > \beta_N M_\ell^{(qN^\xi)}\}$). Recall (6.16) and $N^{-d\xi/\alpha} M_L^{(qN^\xi)} \rightarrow M_L^{(q)}$ in Proposition 4.3). We now prove the analogous of Lemma 5.1 to show that the last term is negligible.

Lemma 6.4. *For any $\varepsilon > 0$ and any $\rho > 0$, we have*

$$\limsup_{L \rightarrow \infty} \limsup_{N \rightarrow +\infty} \mathbb{P} \left(\frac{1}{N^{2\xi-1}} \log Z_{N,\rho\beta_N}^{(L,\ell)} > \varepsilon \right) = 0.$$

Proof. Note that we easily have, as in the calculation leading to (6.18),

$$Z_{N,\rho\beta_N}^{(L,\ell)} \leq \sum_{\Delta \subset \Upsilon_\ell^{(qN^\xi)}} e^{\rho\beta_N \Omega_{qN^\xi}^{(>L)}} \mathbf{P}(\Delta \subset \mathcal{R}_N) \leq \ell! (2C_1)^\ell \exp \left(C_2 \mathcal{T}_{N,qN^\xi}^{C_\rho\beta_N,(>L)} \right). \quad (6.46)$$

Then

$$N^{-(2\xi-1)} \log Z_{N,\rho\beta_N}^{(L,\ell)} \leq \ell N^{-(2\xi-1)} \log(\ell 2C_1) + C_2 N^{-(2\xi-1)} \mathcal{T}_{N,qN^\xi}^{\frac{\rho\beta_N}{C_2},(>L)}.$$

Note that since $\ell = (q^2 N^{2\xi-1})^{1-\delta}$ with q fixed, the first term goes to 0 as $N \rightarrow \infty$ (recall $\xi > 1/2$). For the second term, using that $\beta_N \sim \beta N^{2\xi-1-d\xi/\alpha}$ we get that for any $\varepsilon > 0$

$$\mathbb{P} \left(\mathcal{T}_{N,qN^\xi}^{C_\rho\beta_N,(>L)} > (2C_2)^{-1} \varepsilon N^{2\xi-1} \right) = \mathbb{P} \left(\mathcal{T}_{N,qN^\xi}^{C_\rho\beta_N,(>L)} > C' \varepsilon L^{2(\frac{1}{\alpha}-\frac{1}{d})} N \times \left(C_\rho \beta_N (qN^\xi)^{\frac{d}{\alpha}-1} L^{\frac{1}{d}-\frac{1}{\alpha}} \right)^2 \right),$$

where the constant C' depends on ρ, q, β . Then, using Proposition 4.1-(4.9), provided that L is large enough so that $C' \varepsilon L^{2(\frac{1}{\alpha}-\frac{1}{d})} > 1$, we get

$$\mathbb{P} \left(C_2 \mathcal{T}_{N,qN^\xi}^{C_\rho\beta_N,(>L)} > \frac{1}{2} \varepsilon N^{2\xi-1} \right) \leq c \left(C' \varepsilon L^{2(\frac{1}{\alpha}-\frac{1}{d})} \right)^{-\frac{\alpha L d}{2(\alpha L + d)}} \leq c_\varepsilon L^{-a},$$

for some exponent $a > 0$. This concludes the proof of Lemma 6.4. \square

Step 2.(ii). Once the number of weights is fixed, we can prove the following convergence in distribution. The proof is identical to that of Proposition 5.2 (replacing $\hat{\text{Ent}}$ with Ent) and is omitted.

Proposition 6.5. *For any positive integer L and any $\rho > 0$, we have the following convergence*

$$\frac{1}{N^{2\xi-1}} \log Z_{N,\rho\beta_N}^{(L)} \xrightarrow{(d)} \mathcal{T}_{\rho\beta,q}^{(L)},$$

where $\mathcal{T}_{\rho,q}^{(L)}$ is defined in (4.14).

Step 2.(iii). As mentioned in Proposition 4.2, we have that $\mathcal{T}_{\rho\beta,q}^{(L)}$ converges to $\mathcal{T}_{\rho\beta,q}$ as $L \rightarrow \infty$. Thanks to the continuity in β , we also have that $\mathcal{T}_{\rho\beta,q}$ converges to $\mathcal{T}_{\rho\beta,q}$ as $\rho \rightarrow 1$. The conclusion of the proof of Proposition 6.3 then follows from combining Proposition 6.5 with Lemma 6.4 (and Step 2.(i)-a), letting $L \rightarrow \infty$ then ρ to 1.

6.3 Step 3: letting $q \rightarrow \infty$

Going back to (6.1), the conclusion of the convergence in Theorem 2.7 simply follows from Step 1 (Proposition 6.1) and Step 2 (Proposition 6.3). Indeed, thanks to Proposition 6.1, we get that if q has been fixed large enough, then with probability $1 - cq^{-\nu}$ we have for N large enough

$$\bar{Z}_{N,\beta_N}(M_N \leq qN^\xi) \leq \bar{Z}_{N,\beta_N} \leq \bar{Z}_{N,\beta_N}(M_N \leq qN^\xi) + 1.$$

Using Proposition 6.3 and noting that $\mathcal{T}_\beta > 0$ a.s. and then letting $q \rightarrow \infty$ concludes the proof of (2.1) (recall that $\mathcal{T}_{\beta,q}$ converges to \mathcal{T}_β as $q \rightarrow \infty$, by monotonicity).

6.4 Transversal fluctuations

Let us first prove that the transversal fluctuations are *at least* of order N^ξ , analogously to what is done in Section 5.2. Notice that for any $\eta > 0$ we have thanks to Proposition 6.3 and (2.1) (using Skorokhod's representation theorem, cf. Remark 2.3)

$$\frac{1}{N^{2\xi-1}} \log \bar{\mathbf{P}}_{N,\beta_N}(M_N \leq \eta N^\xi) = \frac{1}{N^{2\xi-1}} \left(\log \bar{Z}_{N,\beta_N}(M_N \leq \eta N^\xi) - \log \bar{Z}_{N,\beta_N} \right) \longrightarrow \mathcal{T}_{\beta,\eta} - \mathcal{T}_\beta \quad \text{a.s.} \quad (6.47)$$

Since we have $\mathcal{T}_\beta > 0$ and $\lim_{\eta \downarrow 0} \mathcal{T}_{\beta,\eta} = 0$ a.s., we can choose η small enough so that $\mathcal{T}_{\beta,\eta} - \mathcal{T}_\beta$ is negative with high \mathbb{P} -probability. From (6.47), on the event $\mathcal{T}_{\beta,\eta} - \mathcal{T}_\beta < 0$ we have that $\bar{\mathbf{P}}_N(M_N \leq \eta N^\xi)$ goes to 0: this concludes the proof that transversal fluctuations are at least of order N^ξ .

Then we prove that the transversal fluctuations are *at most* of order N^ξ . We have that

$$\frac{1}{N^{2\xi-1}} \log \bar{\mathbf{P}}_{N,\beta_N} \left(M_N \geq \frac{1}{\eta} N^\xi \right) = \frac{1}{N^{2\xi-1}} \log \bar{Z}_{N,\beta_N} \left(M_N \geq \frac{1}{\eta} N^\xi \right) - \frac{1}{N^{2\xi-1}} \log \bar{Z}_{N,\beta_N}.$$

The second term above converges to \mathcal{T}_β , which is positive a.s.. For the first term, we have that $\bar{Z}_{N,\beta_N}(M_N \geq \eta^{-1} N^\xi) < \exp(-c_1 \eta^{-2} N^{2\xi-1})$ with \mathbb{P} -probability larger than $1 - \eta^\nu$ by Proposition 6.1. Therefore, $\bar{\mathbf{P}}(M_N \geq \eta^{-1} N^\xi)$ vanishes exponentially fast with \mathbb{P} -probability close to 1 for small enough η , which concludes the result.

7 Region C: proof of Theorems 2.12

We will prove the result in the case $\lim_{N \rightarrow \infty} \beta_N N^{\frac{d}{2\alpha}} = \beta \in [0, +\infty)$. We start with the case $\alpha > 2$ before we turn to the case $\alpha < 2$. We again use the notation \bar{Z}_{N,β_N} for $Z_{N,\beta_N}^{\omega, h=\mu}$.

7.1 The case $\alpha \in (2 \vee \frac{d}{2}, d)$

We first focus on the convergence (2.7) of Theorem 2.12. For simplicity, we treat the case where β_N decays at most polynomially; the case where β_N decays faster (and thus $\beta = 0$) is even simpler. Note that the condition $\alpha \in (2 \vee \frac{d}{2}, d)$ implies in particular that $d \geq 3$, and recall the definition (2.5) of \mathcal{X} in dimension $d \geq 5$.

To simplify the notation, let us suppose that $\mathbb{V}\text{ar}(\omega_x) = 1$ and observe that the convergence (2.7) is equivalent to

$$\frac{1}{a_N \beta_N} \left(e^{-\frac{1}{2} \beta_N^2 \mathbf{E}[\|\mathcal{R}_N\|]} \bar{Z}_{N,\beta_N} - 1 \right) \xrightarrow{(d)} \begin{cases} \mathcal{N}(0, \sigma_d^2) & \text{if } d = 3, 4, \\ \mathcal{X} & \text{if } d \geq 5. \end{cases} \quad (7.1)$$

since both imply that $e^{-\frac{1}{2} \beta_N^2 \mathbf{E}[\|\mathcal{R}_N\|]} \bar{Z}_{N,\beta_N}$ converges to 1 in probability.

Step 1. Truncation of the environment. Let us set

$$k_N := (\log N)^\eta N^{\frac{d}{2\alpha}} \quad (7.2)$$

where $\eta \in (\frac{d}{2\alpha}, 1)$ and let us define the truncated environment

$$\tilde{\omega} = (\omega - \mu) \mathbb{1}_{\{\omega \leq k_n\}}. \quad (7.3)$$

Let us also define $\lambda_N = \log \mathbb{E}[e^{\beta_N \tilde{\omega}_0}]$. The next lemma compares the partition function \bar{Z}_{N, β_N} with its counterpart with the truncated environment $\tilde{\omega}$.

Lemma 7.1. *Let $\alpha \in (2 \vee \frac{d}{2}, d)$. Assume that $\beta_N \leq cN^{-d/2\alpha}$ for some constant c . Then for any $a > 0$, we have that $N^a(\bar{Z}_{N, \beta_N} - Z_{N, \beta_N}^{\tilde{\omega}}) \rightarrow 0$ in \mathbb{P} -probability.*

Proof. First, we show that we can reduce the partition function to trajectories with transversal fluctuations at most $\sqrt{N \log N}$. Assume that $\beta_N N^{d/2\alpha} \leq c$, with $\alpha > d/2$, then for any $A_N \geq 1$

$$\mathbb{P}\left(\bar{Z}_{N, \beta_N}(M_N \geq A_N \sqrt{N \log N}) \geq e^{-c_1 A_N^2 \log N}\right) \leq c_2 A_N^{-\nu}. \quad (7.4)$$

This result is identical to Proposition 6.1, replacing N^ξ with $\sqrt{N \log N}$ (so $N^{2\xi-1}$ is replaced with $\log N$): the proof is identical and relies on Lemma 6.2 (in which we can take $h_N = \sqrt{N \log N}$), so we omit it. Thanks to (7.4), we can therefore choose a sequence A_N going to infinity (let us take $A_N = \log \log N$) in order to get that for any $a > 0$, $N^a Z_{N, \beta_N}^{\omega, h=\mu}(M_N \geq A_N \sqrt{N \log N})$ goes to 0 in probability.

On the other hand, using Cauchy-Swartz inequality we also have that

$$\begin{aligned} \mathbb{E}[Z_{N, \beta_N}^{\tilde{\omega}}(M_N > A_N \sqrt{N \log N})] &= \mathbb{E}[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{M_N > A_N \sqrt{N \log N}\}}] \\ &\leq \mathbb{E}[e^{2\lambda_N |\mathcal{R}_N|}]^{1/2} \mathbf{P}(M_N > A_N \sqrt{N \log N})^{1/2}. \end{aligned} \quad (7.5)$$

Now, since $\alpha > 2$, $\lambda_N \sim \frac{1}{2} \beta_N^2$ (see Lemma C.1). In particular, since $\beta_N \leq cN^{-d/2\alpha}$, we get that $|\lambda_N| \leq C' N^{-d/\alpha}$, so $\lambda_N |\mathcal{R}_N| \leq \lambda_N N$ goes to 0 as $N \rightarrow \infty$ since $\alpha < d$. Also, the last probability in (7.5) is bounded by $\exp(-c A_N^2 \log N) = N^{-c A_N^2}$. By Markov's inequality, we therefore get that $N^a Z_{N, \beta_N}^{\tilde{\omega}}(M_N > A_N \sqrt{N \log N})$ goes to 0 in probability, using again that $A_N \rightarrow \infty$.

All together, it remains to prove that

$$N^a \left(\bar{Z}_{N, \beta_N}(M_N \leq A_N \sqrt{N \log N}) - Z_{N, \beta_N}^{\tilde{\omega}}(M_N \leq A_N \sqrt{N \log N}) \right) \rightarrow 0 \quad \text{in } \mathbb{P}\text{-probability.} \quad (7.6)$$

But a union bound simply gives that

$$\begin{aligned} \mathbb{P}(\exists x, \|x\| \leq A_N \sqrt{N \log N}, \tilde{\omega}_x \neq \omega_x - \mu) &\leq A_N^d (N \log N)^{d/2} \mathbb{P}(\omega > k_N) \\ &\leq c' A_N^d (N \log N)^{d/2} (\log N)^{-\alpha \eta} N^{-d/2}. \end{aligned} \quad (7.7)$$

where we used that $k_N = (\log N)^\eta N^{d/2\alpha}$ and our assumption (1.4). Since $\eta > d/2\alpha$, choosing $A_N = \log \log N$ we have that this probability goes to 0. This proves (7.6) and concludes the proof of Lemma 7.1 \square

Step 2. Convergence of the partition function with truncated environment.

To conclude the proof of Theorem 2.12-(2.7), and in view of (7.1) and Lemma 7.1 (recall we assumed that β_N decays at most polynomially), we simply need to show the following lemma. For simplicity, let us assume that $\text{Var}(\omega) = 1$ and let us denote $\Xi_N := \frac{1}{2} \text{Var}(\omega) \beta_N^2 \mathbb{E}[|\mathcal{R}_N|]$ the centering term in Theorem 2.12-(2.7); note that (cf. Remark 2.13) $\lim_{N \rightarrow \infty} e^{\Xi_N} = 1$. Recall the definition (2.5) of \mathcal{X} , and recall that we have set $a_N = N^{1/4}$ for $d = 3$, $a_N = (\log N)^{1/2}$ for $d = 4$ and $a_N = 1$ for $d \geq 5$.

Lemma 7.2. Assume that $\beta_N \leq cN^{-d/2\alpha}$. If $\alpha \in (2 \vee \frac{d}{2}, d)$, then we have the following convergence in distribution

$$\frac{1}{a_N \beta_N} (Z_{N, \beta_N}^{\tilde{\omega}} - e^{\Xi_N}) \xrightarrow{(d)} \begin{cases} \mathcal{N}(0, \sigma_d^2) & \text{if } d = 3, 4, \\ \mathcal{X} & \text{if } d \geq 5. \end{cases} \quad (7.8)$$

If $d \geq 5$ and $\alpha \in (\frac{d}{d-2}, \frac{d}{2})$, then we have $\frac{1}{\beta_N} (Z_{N, \beta_N}^{\tilde{\omega}} - 1) \xrightarrow{(d)} \mathcal{X}$.

Remark 7.3. We included here the result for $\alpha \in (\frac{d}{d-2}, \frac{d}{2})$ since the proof is identical to that of (7.8); this will be useful to treat the case $\alpha < d/2$.

As a preliminary to the proof of Lemma 7.2, let us collect some important estimates. Thanks to Lemma C.1, we get that $|e^{\lambda_N} - 1| \leq C\beta_N^2$ and $\lambda_N \sim \frac{1}{2}\beta_N^2$ when $\alpha > 2$ and $|e^{\lambda_N} - 1| \leq Ce^{c(\log N)^\eta} \beta_N^\alpha$ when $\alpha \in (1, 2]$. In particular, we have

$$\lambda_N \leq C\beta_N^2 \text{ if } \alpha \in (2, d), \quad \lambda_N \leq Ce^{\beta_N k_N} \beta_N^\alpha \text{ if } \alpha \in (1, 2]. \quad (7.9)$$

Note that (7.9) implies that $\lim_{N \rightarrow \infty} \lambda_N N = 0$.

Also, define $\varepsilon_x := \exp(\beta_N \tilde{\omega}_x - \lambda_N) - 1$, which will be used in both lemmas. Note that $\mathbb{E}[\varepsilon_x] = 0$ and that thanks to Lemma C.1, we get as for λ_N

$$\mathbb{E}[(\varepsilon_0)^2] \leq C\beta_N^2 \text{ if } \alpha \in (2, d), \quad \mathbb{E}[(\varepsilon_0)^2] \leq Ce^{\beta_N k_N} \beta_N^\alpha \text{ if } \alpha \in (1, 2]. \quad (7.10)$$

In fact, we have that $\mathbb{E}[(\varepsilon_0)^2] \sim \beta_N^2$ when $\alpha > 2$. The bounds (7.9)-(7.10), combined with the fact that $\beta_N k_N \leq c(\log N)^\eta$, will be used extensively throughout the proof.

Proof of Lemma 7.2. Note that $e^{u\mathbb{1}_{\{x \in \mathcal{R}_N\}}} = 1 + (e^u - 1)\mathbb{1}_{\{x \in \mathcal{R}_N\}}$ for any $u \in \mathbb{R}$. Writing this with $u = \beta_N \tilde{\omega}_x - \lambda_N$ so that $e^u - 1 = \varepsilon_x$, we expand the product in the partition function as follows:

$$Z_{N, \beta_N}^{\tilde{\omega}} = \mathbf{E} \left[e^{\lambda_N |\mathcal{R}_N|} \prod_{x \in \mathbb{Z}^d} (1 + \varepsilon_x \mathbb{1}_{\{x \in \mathcal{R}_N\}}) \right] = \mathbf{E} [e^{\lambda_N |\mathcal{R}_N|}] + \sum_{x \in \mathbb{Z}^d} \varepsilon_x \mathbf{E} [e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{x \in \mathcal{R}_N\}}] + \mathbf{Y}_N, \quad (7.11)$$

where we set

$$\mathbf{Y}_N := \sum_{k \geq 2} \sum_{\substack{(x_1, \dots, x_k) \in (\mathbb{Z}^d)^k \\ x_i \neq x_j, i, j=1, \dots, k}} \left(\prod_{i=1}^k \varepsilon_{x_i} \right) \mathbf{E} \left[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{(x_1, \dots, x_k) \in \mathcal{R}_N\}} \right], \quad (7.12)$$

where $(x_1, \dots, x_k) \in \mathcal{R}_N$ means that the points x_1, \dots, x_k are visited in this given order; in particular, we do not have the combinatorial term $k!$.

We now control all terms, starting with \mathbf{Y}_N .

Last term in (7.11). We show that $\frac{1}{\beta_N a_N} \mathbf{Y}_N$ converges in probability to zero. Since $\mathbb{E}[\varepsilon_x] = 0$, we only need to control the second moment of \mathbf{Y}_N . Because the ε_x are independent, we easily get that

$$\mathbb{E}[(\mathbf{Y}_N)^2] = \sum_{k \geq 2} \sum_{\substack{(x_1, \dots, x_k) \in (\mathbb{Z}^d)^k \\ x_i \neq x_j, i, j=1, \dots, k}} \mathbb{E}[(\varepsilon_0)^2]^k \mathbf{E} \left[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{(x_1, \dots, x_k) \in \mathcal{R}_N\}} \right]^2. \quad (7.13)$$

Now, using the fact that $|\mathcal{R}_N| \leq N$ and Markov's property, we get

$$\begin{aligned} \sum_{\substack{(x_1, \dots, x_k) \in (\mathbb{Z}^d)^k \\ x_i \neq x_j, i, j=1, \dots, k}} \mathbf{E} \left[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{(x_1, \dots, x_k) \in \mathcal{R}_N\}} \right]^2 &\leq e^{2N\lambda_N} \sum_{\substack{(x_1, \dots, x_k) \in (\mathbb{Z}^d)^k \\ x_i \neq x_j, i, j=1, \dots, k}} \mathbf{P}((x_1, \dots, x_k) \in \mathcal{R}_N)^2 \\ &\leq C \left(\sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_N)^2 \right)^k, \end{aligned}$$

where we also used that $\lambda_N N$ goes to 0, as seen above. We denote $J_N := \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_N)^2$. Then, we have $J_N \sim c_d a_N^2$ as $N \rightarrow \infty$, as outlined in Lemma B.5 in the appendix. Using the bound (7.10) on $\mathbb{E}[(\varepsilon_0)^2]$, we conclude that by $\mathbb{E}[(\varepsilon_x)^2] J_N \rightarrow 0$,

$$\frac{1}{\beta_N^2 a_N^2} \mathbb{E}[(\mathbf{Y}_N)^2] \leq \frac{C}{\beta_N^2 a_N^2} \sum_{k \geq 2} (\mathbb{E}[(\varepsilon_0)^2] J_N)^k \leq C \begin{cases} \beta_N^2 J_N & \text{if } \alpha \in (2, d), \\ e^{2c(\log N)^\eta} \beta_N^{2(\alpha-1)} J_N & \text{if } \alpha \in (1, 2]. \end{cases} \quad (7.14)$$

This readily goes to 0 as $N \rightarrow \infty$.

First term in (7.11). We now show that

$$\lim_{N \rightarrow \infty} \frac{1}{a_N \beta_N} \left| \mathbf{E}[e^{\lambda_N \mathcal{R}_N}] - e^{\Xi_N \mathbb{1}_{\{\alpha > 2\}}} \right| = 0. \quad (7.15)$$

Let us start with the case $\alpha > 2$. Lemma C.1 gives in that case that $|\lambda_N - \frac{1}{2}\beta_N^2| \leq C e^{c'(\log N)^\eta} \beta_N^{3 \wedge \alpha}$. Hence, we get

$$\frac{1}{\beta_N a_N} \left| \mathbf{E}[e^{\lambda_N |\mathcal{R}_N|}] - \mathbf{E}[e^{\frac{1}{2}\beta_N^2 |\mathcal{R}_N|}] \right| \leq \frac{1}{\beta_N a_N} |\lambda_N - \frac{1}{2}\beta_N^2| N \leq C e^{c'(\log N)^\eta} \frac{1}{a_N} \beta_N^{2 \wedge (\alpha-1)} N,$$

which goes to 0 as $N \rightarrow \infty$. Indeed, if $\alpha \in (2, 3)$, we have that $\beta_N^{\alpha-1} \leq cN^{-\frac{d(\alpha-1)}{2\alpha}}$: in dimension $d \geq 4$ we get that $\beta_N^{\alpha-1} N \leq cN^{1-2(\alpha-1)/\alpha}$ which goes to 0 since $\alpha > 2$; in dimension $d = 3$, using that $a_N = N^{1/4}$ we get that $a_N^{-1} \beta_N^{\alpha-1} N \leq N^{\frac{3}{4}(1-2(\alpha-1)/\alpha)}$ which also goes to 0 since $\alpha > 2$. If $\alpha \in [3, d)$, we simply use that $\beta_N^2 N \leq CN^{1-d/\alpha}$, which goes to 0 since $\alpha < d$. Additionnally, recalling that $\Xi_N = \frac{1}{2}\beta_N^2 \mathbf{E}[|\mathcal{R}_N|]$, we have

$$\begin{aligned} \frac{1}{\beta_N a_N} \left| \mathbf{E}[e^{\frac{1}{2}\beta_N^2 |\mathcal{R}_N|}] - e^{\Xi_N} \right| &\leq \frac{C}{\beta_N a_N} e^{\frac{1}{2}\beta_N^2 \mathbf{E}[|\mathcal{R}_N|]} \times \beta_N^4 \mathbf{Var}(|\mathcal{R}_N|) \\ &\leq C' \beta_N^3 N \log N, \end{aligned}$$

where we expanded $e^{\frac{1}{2}\beta_N^2 (|\mathcal{R}_N| - \mathbf{E}[|\mathcal{R}_N|])} - 1$ to get the first inequality (recall that $\beta_N^2 N$ goes to 0), and used that $\mathbf{Var}(|\mathcal{R}_N|) \sim c_3 N \log N$ in dimension $d = 3$ and $\mathbf{Var}(|\mathcal{R}_N|) \sim c_4 N$ in dimension $d \geq 4$ (see [26]), together with the definition of a_N . This proves (7.15) in the case $\alpha \in (2, d)$.

In the case $\alpha \in (\frac{d}{d-2}, 2]$ (in particular $d \geq 5$), we use (7.9) to get that

$$\frac{1}{\beta_N} \left| \mathbf{E}[e^{\lambda_N |\mathcal{R}_N|}] - 1 \right| \leq c \beta_N^{-1} \lambda_N N \leq C e^{c(\log N)^\eta} \beta_N^{\alpha-1} N,$$

which goes to 0 as $N \rightarrow \infty$, because $\frac{d}{2\alpha}(\alpha-1) > 1$. This proves (7.15) in the case $\alpha \in (\frac{d}{d-2}, 2]$.

Second term in (7.11). Let us rewrite it as

$$\frac{1}{a_N} \sum_{x \in \mathbb{Z}^d} \frac{\varepsilon_x}{\beta_N} \mathbf{P}(x \in \mathcal{R}_N) + \frac{1}{\beta_N a_N} \sum_{x \in \mathbb{Z}^d} \varepsilon_x \mathbf{E}[(e^{\lambda_N |\mathcal{R}_N|} - 1) \mathbb{1}_{\{x \in \mathcal{R}_N\}}]. \quad (7.16)$$

The second term in (7.16) goes to zero in probability. Indeed, it is centered in expectation, and its second moment is bounded by

$$\frac{1}{\beta_N^2 a_N^2} \sum_{x \in \mathbb{Z}^d} \mathbb{E}[(\varepsilon_0)^2] (e^{\lambda_N N} - 1)^2 \mathbf{P}(x \in \mathcal{R}_N)^2 \leq C \frac{\mathbb{E}[(\varepsilon_0)^2] (\lambda_N N)^2 J_N}{\beta_N^2 a_N^2}.$$

Using that $J_N \sim c_d a_N^2$ as $N \rightarrow \infty$ (cf. Lemma B.5) and that $\lambda_N, \mathbb{E}[(\varepsilon_0)^2] \leq C e^{c(\log N)^\eta} \beta_N^{\alpha \wedge 2}$, see (7.9) and (7.10), this is bounded by a constant times $e^{c(\log N)^\eta} \beta_N^{3(\alpha \wedge 2) - 2} N^2$. If $\alpha > 2$

this is bounded by $N^{o(1)}(\beta_N^2 N)^2$ so this goes to 0; if $\alpha \in (\frac{d}{d-2}, 2]$ this is bounded by $N^{o(1)}(\beta_N^{\alpha-1} N)^2$, so this also goes to 0, as seen above.

We now rewrite the first term in (7.16) as

$$\frac{1}{a_N} \sum_{x \in \mathbb{Z}^d} (\omega_x - \mu) \mathbf{P}(x \in \mathcal{R}_N) + \frac{1}{a_N \beta_N} \sum_{x \in \mathbb{Z}^d} (\varepsilon_x - \beta_N(\omega_x - \mu)) \mathbf{P}(x \in \mathcal{R}_N), \quad (7.17)$$

and we now show that the second term goes to 0 in probability.

If $\alpha \in (\frac{d}{d-2}, 2]$ ($d \geq 5$), we control the expectation of its absolute value. Using that $|\beta_N \tilde{\omega}_x| \leq \beta_N k_N \leq c(\log N)^\eta$ and recalling that $\varepsilon_x = e^{\beta_N \tilde{\omega}_x - \lambda_N} - 1$, we can perform a Taylor expansion to get that

$$\begin{aligned} \frac{1}{\beta_N} \mathbb{E}[|\varepsilon_x - \beta_N(\omega_x - \mu)|] &\leq \frac{1}{\beta_N} \left(\beta_N \mathbb{E}[|\omega_x - \mu| \mathbb{1}_{\{\omega_x > k_N\}}] + \lambda_N + C e^{2c(\log N)^\eta} (\beta_N^2 \mathbb{E}[\tilde{\omega}_x^2] + \lambda_N^2) \right) \\ &\leq k_N^{1-\alpha} + e^{c'(\log N)^\eta} \beta_N^{\alpha-1}, \end{aligned}$$

where we used that $\mathbb{E}[\tilde{\omega}_x^2] \leq C k_N^{2-\alpha} \log k_N \leq C' \beta_N^{\alpha-2} (\log N)^C$, together with a similar bound on λ_N , see (7.9). Therefore, the expectation of the absolute value of the second term in (7.17) is bounded by a constant times $k_N^{1-\alpha} N + e^{c'(\log N)^\eta} \beta_N^{\alpha-1} N$: this goes to 0 as $N \rightarrow \infty$, thanks to the condition $\alpha > d/(d-2)$. By Markov's inequality, the second term in (7.17) goes to 0 in probability.

If $\alpha \in (2, d)$, since the expectation of the second term in (7.17) is 0, we control its second moment: by independence of the ω_x , it is equal to

$$\frac{1}{a_N^2 \beta_N^2} \sum_{x \in \mathbb{Z}^d} \mathbb{E}[(\varepsilon_x - \beta_N(\omega_x - \mu))^2] \mathbf{P}(x \in \mathcal{R}_N)^2 \leq \frac{C}{\beta_N^2} \mathbb{E}[(\varepsilon_x - \beta_N(\omega_x - \mu))^2],$$

where we have used that $J_N = \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_N) \leq c a_N^2$, see Lemma B.5. Now, we can use the same Taylor expansion as above, to get that

$$\begin{aligned} \frac{1}{\beta_N^2} \mathbb{E}[(\varepsilon_x - \beta_N(\omega_x - \mu))^2] &\leq \frac{C}{\beta_N^2} \left(\beta_N^2 \mathbb{E}[(\omega_x - \mu)^2 \mathbb{1}_{\{\omega_x > k_N\}}] + \lambda_N^2 + C e^{4c(\log N)^\eta} (\beta_N^4 \mathbb{E}[\tilde{\omega}_x^4] + \lambda_N^4) \right) \\ &\leq C k_N^{2-\alpha} + C e^{c'(\log N)^\eta} \beta_N^{\alpha \wedge 4 - 2}. \end{aligned}$$

where we used here that $\lambda_N \sim \frac{1}{2} \beta_N^2$ as $N \rightarrow \infty$ and that $\mathbb{E}[\tilde{\omega}_x^4]$ is bounded by $C k_N^{4-\alpha} \log k_N \leq C \beta_N^{\alpha-4} (\log N)^{C'}$ if $\alpha \leq 4$ and by a constant if $\alpha > 4$. This shows that the variance of the second term in (7.17) goes to 0, so this term goes to 0 in probability.

All together, combining (7.11) with (7.14)-(7.15) and (7.16)-(7.17), we have shown that

$$\frac{1}{a_N \beta_N} \left(Z_{N, \beta_N}^{\tilde{\omega}} - e^{\Xi_N \mathbb{1}_{\{\alpha > 2\}}} - \sum_{x \in \mathbb{Z}^d} \beta_N(\omega_x - \mu) \mathbf{P}(x \in \mathcal{R}_N) \right) \xrightarrow{\mathbb{P}} 0.$$

It therefore only remains to show that

$$\frac{1}{a_N} \sum_{x \in \mathbb{Z}^d} (\omega_x - \mu) \mathbf{P}(x \in \mathcal{R}_N) \xrightarrow{(d)} \begin{cases} \mathcal{N}(0, \sigma_d^2) & \text{if } d = 3, 4, \\ \mathcal{X} & \text{if } d \geq 5. \end{cases} \quad (7.18)$$

Note that in Lemma 7.2, when $d \geq 5$ and $\alpha \in (2, \frac{d}{2})$ we used a centering equal to 1 instead of e^{Ξ_N} . This is not an issue since in that case we have that $\beta_N^{-1}(e^{\Xi_N} - 1) \leq c \beta_N N$ goes to 0 as $N \rightarrow \infty$.

In dimension $d \geq 5$, (7.18) follows simply by monotone convergence, recalling that \mathcal{X} is a.s. finite if $\alpha > \frac{d}{d-2}$, see Proposition 2.10.

In dimension $d = 3, 4$, the random variables $a_N^{-1}(\omega_x - \mu) \mathbf{P}(x \in \mathcal{R}_N)$ are independent, centered and have variance $a_N^{-2} \mathbf{P}(x \in \mathcal{R}_N)^2$. Since $a_N \rightarrow \infty$, this goes to 0 uniformly for $x \in \mathbb{Z}^d$. Hence, we directly get (7.18), for instance with a characteristic function analysis with the constant $\sigma_d^2 := \lim_{N \rightarrow \infty} a_N^{-2} \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_N)^2$, recall Lemma B.5. \square

7.2 The case $\alpha \in (\frac{d}{2}, 2)$

Note that this regime is nonempty only for $d = 2, 3$. Also, since $\alpha > \frac{d}{2} \geq 1$, we have $\mu := \mathbb{E}[\omega_0] < +\infty$. Let us stress that, similarly as (7.1) above, the convergence (2.8) is equivalent to the convergence in distribution

$$\frac{v_N}{\beta_N N^{\frac{d}{2\alpha}}} (\bar{Z}_{N,\beta_N} - 1) \xrightarrow{(d)} \mathcal{W}_\beta. \quad (7.19)$$

The steps are the same as in Section 7.1. The case of the dimension $d = 3$ is very similar, but some adaptations are needed in dimension $d = 2$.

Step 1. Truncation of the environment. Let us set

$$k_N := \begin{cases} (\log N)^\eta N^{\frac{d}{2\alpha}} & \text{if } d = 3, \\ (\log \log N)^\eta N^{\frac{d}{2\alpha}} & \text{if } d = 2, \end{cases} \quad (7.20)$$

where $\eta \in (\frac{d}{2\alpha}, 1)$. We again define the truncated environment $\tilde{\omega} = (\omega - \mu) \mathbb{1}_{\{\omega \leq k_n\}}$ and we set $\lambda_N = \log \mathbb{E}[e^{\beta_N \tilde{\omega}_0}]$. The next lemma is analogous to Lemma 7.1.

Lemma 7.4. *Let $\alpha \in (2, \frac{d}{2})$. Assume that $\beta_N \leq cN^{-d/2\alpha}$ for some constant c . Then for any $a > 0$, we have that:*

- (i) if $d = 3$, $N^a (\bar{Z}_{N,\beta_N} - Z_{N,\beta_N}^{\tilde{\omega}}) \rightarrow 0$ in \mathbb{P} -probability;
- (ii) if $d = 2$, $(\log N)^a (\bar{Z}_{N,\beta_N} - Z_{N,\beta_N}^{\tilde{\omega}}) \rightarrow 0$ in \mathbb{P} -probability.

Before we start the proof, let us stress that Lemma C.1 gives some estimates on λ_N : since $\alpha \in (1, 2)$, we have

$$\lambda_N \leq C e^{\beta_N k_N} \beta_N^\alpha \leq C e^{\beta_N k_N} N^{-d/2}, \quad \text{with } \beta_N k_N = \begin{cases} (\log N)^\eta & \text{if } d = 3, \\ (\log \log N)^\eta & \text{if } d = 2. \end{cases} \quad (7.21)$$

Proof. In the case of dimension $d = 3$, the proof is identical to that of Lemma 7.1: the only difference here is that we do not have $\lambda_N \leq C\beta_N^2$. Instead, we use (7.21) instead which gives that $\lambda_N N \leq C e^{(\log N)^\eta} N^{1-\frac{d}{2}}$ goes to 0 as $N \rightarrow \infty$; this was used to bound (7.5).

We therefore focus on the case of dimension $d = 2$; the idea of the proof is identical, with some adaptation. First of all, analogously to (7.4), we get that for any sequence $A_N \geq 1$

$$\mathbb{P} \left(\bar{Z}_{N,\beta_N} (M_N \geq A_N \sqrt{N \log \log N}) \geq e^{-c_1 A_N^2 \log \log N} = (\log N)^{-c_1 A_N^2} \right) \leq c_2 A_N^{-\nu}. \quad (7.22)$$

Again, this is identical to Proposition 6.1, replacing N^ξ with $\sqrt{N \log \log N}$: in particular, this relies on Lemma 6.2 in which we can take $h_N = \sqrt{N \log N \log N}$. Therefore, choosing a sequence $A_N \rightarrow \infty$ (we take $A_N = \log \log N$), we get that $(\log N)^a Z_{N,\beta_N}^{\omega, h=\mu} (M_N \geq A_N \sqrt{N \log \log N})$ goes to 0 in probability.

Then, as in (7.5), we have that

$$\begin{aligned} \mathbb{E} [Z_{N,\beta_N}^{\tilde{\omega}} (M_N > A_N \sqrt{N \log \log N})] &\leq \mathbb{E} [e^{2\lambda_N |\mathcal{R}_N|}]^{1/2} \mathbb{P} (M_N > A_N \sqrt{N \log \log N})^{1/2} \\ &\leq \mathbb{E} [e^{2\lambda_N |\mathcal{R}_N|}]^{1/2} e^{-c A_N^2 \log \log N}. \end{aligned} \quad (7.23)$$

The main difference here is that we cannot use the bound $\lambda_N |\mathcal{R}_N| \leq \lambda_N N$ since we do not have $\lambda_N N \rightarrow 0$ as $N \rightarrow \infty$, see (7.21). However, we may use the fact that $|\mathcal{R}_N|$ is of the order $N/\log N$ to obtain that $\lambda_N |\mathcal{R}_N| \rightarrow 0$ with high \mathbb{P} -probability. More precisely, using that $\mathbb{E}[|\mathcal{R}_N|] \leq CN/\log N$ (see e.g. [16, Eq. (5.3.39)]) we have that, for a sufficiently large C'

$$\mathbb{P} \left(|\mathcal{R}_N| \geq 2C' \frac{N}{\log N} \right) \leq \mathbb{P} \left(|\mathcal{R}_N| - \mathbb{E}[|\mathcal{R}_N|] \geq C' \frac{N \log \log N}{(\log N)^2} \right) \leq e^{-c(\log N)^4}. \quad (7.24)$$

where we used [16, Thm. 8.5.1] for the last inequality. We therefore get that

$$\mathbf{E}[e^{2\lambda_N |\mathcal{R}_N|}] \leq e^{4C' \lambda_N N / \log N} + e^{2\lambda_N N} e^{-c(\log N)^4}.$$

Now, thanks to (7.21), we get that $\lambda_N N = o(\log N)$ as $N \rightarrow \infty$. Hence, from (7.23), we get that

$$\mathbf{E}[Z_{N,\beta_N}^{\tilde{\omega}}(M_N > A_N \sqrt{N \log \log N})] \leq C(\log N)^{-cA_N^2}$$

so that $(\log N)^a Z_{N,\beta_N}^{\tilde{\omega}}(M_N > A_N \sqrt{N \log \log N}) \rightarrow 0$ in \mathbb{P} -probability thanks to Markov's inequality.

Finally, notice that similarly to (7.7), we have that

$$\mathbb{P}(\exists x, \|x\| \leq A_N \sqrt{N \log \log N}, \tilde{\omega}_x \neq \omega_x - \mu) \leq c' A_N^d (N \log \log N)^{d/2} (\log \log N)^{-\alpha \eta} N^{-d/2}, \quad (7.25)$$

which goes to 0 as $N \rightarrow \infty$ since $\alpha \eta > d/2$, and $A_N = \log \log \log N$. Therefore, we get that $Z_{N,\beta_N}^{\tilde{\omega}}(M_N > A_N \sqrt{N \log \log N})$ is equal to $Z_{N,\beta_N}^{\tilde{\omega}}(M_N > A_N \sqrt{N \log \log N})$ with \mathbb{P} -probability going to 1. This concludes the proof of Lemma 7.4. \square

Step 2. Convergence of the partition function with truncated environment.

To conclude the proof of Theorem 2.12–(2.8), and in view of (7.19) and Lemma 7.4, it remains to prove the following lemma. (In dimension $d = 2$, we need to assume that $\beta_N N^{d/2\alpha}$ decays slower than any power of $\log N$, but this does not hide anything deep; one simply needs to use a more restrictive truncation, but this actually simplifies many of the arguments so we do not treat this case.)

Lemma 7.5. *Let $\alpha \in (\frac{d}{2}, 2)$. If $\lim_{N \rightarrow \infty} \beta_N N^{d/2\alpha} = \beta \in [0, +\infty)$ then we have the following convergence in distribution*

$$\frac{v_N}{\beta_N N^{d/2\alpha}} (Z_{N,\beta_N}^{\tilde{\omega}} - 1) \xrightarrow{(d)} \mathcal{W}_\beta.$$

Proof. We use the same notation as in Step 2 of Section 7.1. Defining $\varepsilon_x := \exp(\beta_N \tilde{\omega}_x - \lambda_N) - 1$, and using a polynomial chaos expansion similar to (7.11), we get

$$Z_{N,\beta_N}^{\tilde{\omega}} - 1 = \mathbf{E}[e^{\lambda_N |\mathcal{R}_N|} - 1] + \sum_{x \in \mathbb{Z}^d} \varepsilon_x \mathbf{E}[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{x \in \mathcal{R}_N\}}] + \mathbf{Y}_N. \quad (7.26)$$

where we defined \mathbf{Y}_N as in (7.12). We first prove that the last term is negligible: the computation in $d = 3$ follows closely the computations (7.13)–(7.14), while dimension $d = 2$ is more delicate. In the second part of the proof we show that the first two terms give the main contribution to the convergence. This part is quite technical and we split it into three steps.

Last term in (7.26). Let us prove that $v_N \mathbf{Y}_N / (\beta_N N^{d/2\alpha})$ goes to 0 in probability.

We start with the case of dimension $d = 3$, which we essentially already treated. Reasoning as in (7.13)–(7.14), we computed the second moment of \mathbf{Y}_N in the case $\alpha \in (1, 2)$: we have that

$$\mathbf{E}[\mathbf{Y}_N^2] \leq C \sum_{k \geq 2} (\mathbf{E}[\varepsilon_0^2] J_N)^k \leq C e^{2c(\log N)^\eta} \beta_N^{2\alpha} N, \quad (7.27)$$

where we bounded $|\mathcal{R}_N| \leq N$, used that $J_N \leq cN^{1/2}$ in dimension $d = 3$ (see Lemma B.5) and the bound (7.10). Since $v_N = N^{1/2}$, we obtain

$$\left(\frac{v_N}{\beta_N N^{d/2\alpha}} \right)^2 \mathbf{E}[\mathbf{Y}_N^2] \leq C e^{2(\log N)^\eta} N^{1-\frac{d}{\alpha}} \beta_N^{2(\alpha-1)} \leq C' e^{c(\log N)^\eta} N^{2-d}. \quad (7.28)$$

This term goes to 0, which concludes the proof in the case of dimension $d = 3$.

Turning to the case of dimension $d = 2$, we need to be more careful: we cannot simply bound $|\mathcal{R}_N|$ by N since we do not have $\lambda_N N \rightarrow 0$ anymore. The calculation of $\mathbb{E}[\mathbf{Y}_N^2]$ in (7.13) remains valid. Then, decomposing the expectation according to whether $|\mathcal{R}_N| \leq 2C'N/\log N$ or not, we obtain

$$\begin{aligned} & \mathbb{E} \left[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{(x_1, \dots, x_k) \in \mathcal{R}_N\}} \right]^2 \\ & \leq \left(e^{2C' \frac{\lambda_N N}{\log N}} \mathbf{P}((x_1, \dots, x_k) \in \mathcal{R}_N) + e^{\lambda_N N} \mathbf{P}((x_1, \dots, x_k) \in \mathcal{R}_N; |\mathcal{R}_N| \geq \frac{2C'N}{\log N}) \right)^2 \\ & \leq 2e^{4C' \frac{\lambda_N N}{\log N}} \mathbf{P}((x_1, \dots, x_k) \in \mathcal{R}_N)^2 + 2e^{2\lambda_N N} \mathbf{P}((x_1, \dots, x_k) \in \mathcal{R}_N) \mathbf{P}(|\mathcal{R}_N| \geq \frac{2C'N}{\log N}), \end{aligned} \quad (7.29)$$

where for the last inequality we used that $(a+b)^2 \leq 2(a^2 + b^2)$ together with Cauchy-Schwarz inequality. Now, we have that $\lambda_N N = o(\log N)$ (recall (7.21)) and also that $\mathbf{P}(|\mathcal{R}_N| \geq 2C'N/\log N) \leq e^{-c(\log N)^4}$ if C' has been fixed large enough, thanks to (7.24). Hence, summing (7.29) over (x_1, \dots, x_k) and using Markov's property, we finally get that

$$\sum_{\substack{(x_1, \dots, x_k) \in (\mathbb{Z}^d)^k \\ x_i \neq x_j, i, j=1, \dots, k}} \mathbb{E} \left[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{(x_1, \dots, x_k) \in \mathcal{R}_N\}} \right]^2 \leq C \left(\sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_N)^2 \right)^k + e^{-c'(\log N)^4} \left(\sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_N) \right)^k$$

All together, going back to (7.13), we obtain that when $d = 2$ the bound (7.27) is replaced with

$$\mathbb{E}[\mathbf{Y}_N^2] \leq C \sum_{k \geq 2} (\mathbb{E}[(\varepsilon_0)^2] J_N)^k + e^{-c'(\log N)^4} \sum_{k \geq 2} (\mathbb{E}[(\varepsilon_0)^2] \mathbb{E}[|\mathcal{R}_N|])^k.$$

Now, we have that $\mathbb{E}[(\varepsilon_0)^2] \leq C e^{(\log \log N)^\eta} \beta_N^\alpha$ by Lemma C.1, that $J_N \leq C \frac{N}{(\log N)^2}$ by Lemma B.5 and that $\mathbb{E}[|\mathcal{R}_N|] \leq C \frac{N}{\log N}$ by [16, Eq. (5.3.39)]. Using that $v_N = \log N$, we obtain

$$\left(\frac{v_N}{\beta_N N^{\frac{d}{2\alpha}}} \right)^2 \mathbb{E}[\mathbf{Y}_N^2] \leq C (\log N)^2 e^{2(\log \log N)^\eta} \beta_N^{2(\alpha-1)} N^{2-\frac{d}{\alpha}} \left(\frac{1}{(\log N)^4} + e^{-c'(\log N)^4} \frac{1}{(\log N)^2} \right), \quad (7.30)$$

and note that $\beta_N^{2(\alpha-1)} N^{2-d/\alpha} \leq c$. This term goes to 0, which concludes the proof when $d = 2$.

First terms in (7.26). We can rewrite these terms as

$$\begin{aligned} & \mathbb{E}[e^{\lambda_N |\mathcal{R}_N|} - 1] + \sum_{x \in \mathbb{Z}^d} \varepsilon_x \mathbb{E}[(e^{\lambda_N |\mathcal{R}_N|} - 1) \mathbb{1}_{\{x \in \mathcal{R}_N\}}] + \sum_{x \in \mathbb{Z}^d} \varepsilon_x \mathbf{P}(x \in \mathcal{R}_N) \\ & = \mathbb{E}[e^{\lambda_N |\mathcal{R}_N|} - 1 - (1 - e^{-\lambda_N}) |\mathcal{R}_N|] + \sum_{x \in \mathbb{Z}^d} \varepsilon_x \mathbb{E}[(e^{\lambda_N |\mathcal{R}_N|} - 1) \mathbb{1}_{\{x \in \mathcal{R}_N\}}] \\ & \quad + e^{-\lambda_N} \sum_{x \in \mathbb{Z}^d} (e^{\beta_N \tilde{\omega}_x} - 1) \mathbf{P}(x \in \mathcal{R}_N), \end{aligned} \quad (7.31)$$

where we used that $\varepsilon_x = e^{-\lambda_N} (e^{\beta_N \tilde{\omega}_x} - 1) - 1 + e^{-\lambda_N}$ for the last identity. Let us denote **I**, **II** and **III** the three terms in (7.31). We show that the first two terms are negligible.

Term I. For the first term in (7.31), let us first treat the case of dimension $d = 3$, which is simpler. Using that $\lambda_N \rightarrow 0$ and $\lambda_N |\mathcal{R}_N| \leq \lambda_N N \rightarrow 0$, a Taylor expansion gives that it is bounded by a constant times $(\lambda_N N)^2$. Hence, using that $\lambda_N \leq C e^{c(\log N)^\eta} \beta_N^\alpha$, we get that

$$\frac{v_N}{\beta_N N^{\frac{d}{2\alpha}}} \mathbf{I} \leq C' e^{2c(\log N)^\eta} v_N \beta_N^{2\alpha-1} N^{2-\frac{d}{2\alpha}} \leq C' e^{c(\log N)^\eta} N^{\frac{5}{2}-d},$$

where we used that $v_N = N^{1/2}$ in dimension $d = 3$. This goes to 0 as $N \rightarrow \infty$.

The case of dimension $d = 2$ is a bit more delicate. On the event $\{|\mathcal{R}_N| \leq 2C'N/\log N\}$, we can again use a Taylor expansion, since $\lambda_N N/\log N \rightarrow 0$, so that the term inside the expectation is bounded in absolute value by a constant times $\lambda_N^2 N^2/(\log N)^2$. On the event $\{|\mathcal{R}_N| > 2C'N/\log N\}$, we bound the absolute value of the term inside the expectation by $e^{N\lambda_N} - 1 + c\lambda_N N \leq C\lambda_N N e^{o(\log N)}$. We therefore get that

$$|\mathbf{I}| \leq C \frac{\lambda_N^2 N^2}{(\log N)^2} + \lambda_N N e^{o(\log N)} \mathbf{P}\left(|\mathcal{R}_N| > 2C' \frac{N}{\log N}\right) \leq \frac{e^{2(\log \log N)^\eta}}{(\log N)^2} \beta_N^{2\alpha} N^2 + e^{-c'(\log N)^4} \beta_N^\alpha N,$$

where we used (7.24) for the last inequality. Since $v_N = \log N$, we get that

$$\frac{v_N}{\beta_N N^{\frac{d}{2\alpha}}} \mathbf{I} \leq \frac{e^{2(\log \log N)^\eta}}{\log N} \beta_N^{2\alpha-1} N^{2-\frac{d}{2\alpha}} + \log N e^{-c'(\log N)^4} \beta_N^{\alpha-1} N^{1-\frac{d}{2\alpha}} \quad (7.32)$$

which goes to 0 as $N \rightarrow \infty$, noticing that $\beta_N^{\alpha-1} N^{1-d/2\alpha} \leq c$, $\beta_N^{2\alpha-1} N^{2-\frac{d}{2\alpha}} \leq c$ and $\eta \in (\frac{d}{2\alpha}, 1)$.

Term II. We prove that $v_N \mathbf{II}/(\beta_N N^{d/2\alpha})$ goes to 0 in probability, by controlling its second moment. Since the ε_x are centered and independent, we have

$$\mathbb{E}[(\mathbf{II})^2] \leq \mathbb{E}[\varepsilon_0^2] \sum_{x \in \mathbb{Z}^d} \mathbb{E}\left[(e^{\lambda_N |\mathcal{R}_N|} - 1) \mathbb{1}_{\{x \in \mathcal{R}_N\}}\right]^2. \quad (7.33)$$

We start with the simpler case of dimension $d = 3$. Using that $e^{\lambda_N |\mathcal{R}_N|} - 1 \leq c\lambda_N N$ and recalling the definition J_N , we get that

$$\mathbb{E}[(\mathbf{II})^2] \leq c\mathbb{E}[\varepsilon_0^2](\lambda_N N)^2 J_N \leq C e^{3c(\log N)^\eta} \beta_N^{3\alpha} N^{\frac{5}{2}},$$

where we used the bounds (7.9)-(7.10) together with the fact that $J_N \leq N^{1/2}$ in dimension $d = 3$, see Lemma B.5. Hence, using also that $v_N = N^{1/2}$ we get that

$$\left(\frac{v_N}{\beta_N N^{d/2\alpha}}\right)^2 \mathbb{E}[(\mathbf{II})^2] \leq C' e^{c'(\log N)^\eta} \beta_N^{3\alpha-2} N^{\frac{7}{2}-\frac{d}{\alpha}} \leq C' e^{c'(\log N)^\eta} N^{\frac{7}{2}-\frac{3d}{2}},$$

which goes to 0 as $N \rightarrow \infty$.

In the case of dimension $d = 2$, we proceed as above. Writing $e^{\lambda_N |\mathcal{R}_N|} - 1 \leq \lambda_N |\mathcal{R}_N| e^{\lambda_N |\mathcal{R}_N|}$ and decomposing according to whether we have $|\mathcal{R}_N| \leq 2C'N/\log N$ or not, we get similarly to (7.29)

$$\mathbb{E}\left[(e^{\lambda_N |\mathcal{R}_N|} - 1) \mathbb{1}_{\{x \in \mathcal{R}_N\}}\right]^2 \leq C \left(\frac{\lambda_N N}{\log N}\right)^2 \mathbf{P}(x \in \mathcal{R}_N)^2 + \lambda_N N e^{\lambda_N N} \mathbf{P}(x \in \mathcal{R}_N) \mathbf{P}\left(|\mathcal{R}_N| \geq 2C' \frac{N}{\log N}\right).$$

Summing over x and using the bound (7.24) together with $\lambda_N N = o(\log N)$ and $\mathbb{E}[|\mathcal{R}_N|] \leq N$, we get from (7.33) that

$$\begin{aligned} \mathbb{E}[(\mathbf{II})^2] &\leq C\mathbb{E}[\varepsilon_0^2] \frac{(\lambda_N N)^2}{(\log N)^2} J_N + C\mathbb{E}[\varepsilon_0^2] \lambda_N N^2 e^{-c(\log N)^4} \\ &\leq C' e^{3(\log \log N)^\eta} \beta_N^{2\alpha} N^2 \left((\log N)^{-4} + e^{-c'(\log N)^4}\right). \end{aligned}$$

For the second inequality, we used the bounds (7.9)-(7.10) (and also that $\beta_N^\alpha N \leq c$) and that $J_N \leq CN/(\log N)^2$, see Lemma B.5. Since $v_N = \log N$, we get that

$$\left(\frac{v_N}{\beta_N N^{d/2\alpha}}\right)^2 \mathbb{E}[(\mathbf{II})^2] \leq C' e^{3(\log \log N)^\eta} \beta_N^{2(\alpha-1)} N^{2-\frac{d}{\alpha}} \left((\log N)^{-2} + (\log N)^2 e^{-c'(\log N)^4}\right), \quad (7.34)$$

which goes to 0 as $N \rightarrow \infty$, using that $\beta_N^{\alpha-1} N^{1-d/2\alpha} \leq c$

Term III. We show the following lemma to control the third term in (7.31).

Lemma 7.6. *Let $d = 2, 3$ and $\alpha \in (\frac{d}{2}, 2)$. We have the following convergence in distribution:*

$$\frac{v_N}{\beta_N N^{d/2\alpha}} \sum_{x \in \mathbb{Z}^d} (e^{\beta_N \tilde{\omega}_x} - 1) \mathbf{P}(x \in \mathcal{R}_N) \xrightarrow{(d)} \mathcal{W}_\beta, \quad (7.35)$$

with \mathcal{W}_β defined in (2.6).

Going back to (7.26), and combining (7.27)–(7.30) with (7.31) and the estimates (7.32)–(7.34), then Lemma 7.6 concludes the proof of Lemma 7.5, noting also that $e^{-\lambda_N}$ goes to 1 as $N \rightarrow \infty$. \square

Proof of Lemma 7.6. To prove (7.35), we adapt the method developed in [19, Thm. 1.4]. We consider only the case $\lim_{N \rightarrow \infty} \beta_N N^{d/2\alpha} = \beta \in (0, +\infty)$; the case $\beta = 0$ can be dealt with similarly, one simply needs to keep track of the dependence in β in all estimates — note also that we have $\mathcal{W}_\beta \rightarrow \mathcal{W}_0$ as $\beta \rightarrow 0$. Also, for $x \in \mathbb{Z}^d$, we use the notation $f_N(x) = v_N \mathbf{P}(x \in \mathcal{R}_N)$.

We prove the convergence of $\sum_{x \in \mathbb{Z}^d} (\beta_N N^{d/2\alpha})^{-1} (e^{\beta_N \tilde{\omega}_x} - 1) f_N(x)$. As in Equation (39) in [19], we fix $\varepsilon > 0$, $K > 0$, and we define $\tilde{\omega}_x^\varepsilon = (\omega_x - \mu) \mathbb{1}_{\{(\omega_x - \mu) \leq \varepsilon N^{d/2\alpha}\}}$. Recall the definition of $\tilde{\omega}_x$ in (7.20). We use the following decomposition of $\sum_{x \in \mathbb{Z}^d} (\beta_N N^{d/2\alpha})^{-1} (e^{\beta_N \tilde{\omega}_x} - 1) f_N(x)$:

$$\begin{aligned} & \sum_{\substack{\|x\| \leq K\sqrt{N} \\ (\omega_x - \mu) > \varepsilon N^{d/2\alpha}}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \tilde{\omega}_x} - 1) f_N(x) + N^{-d/2\alpha} \mathbb{E}[\tilde{\omega}_0^\varepsilon] \sum_{\|x\| \leq K\sqrt{N}} f_N(x) \\ & + \sum_{\|x\| \leq K\sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \tilde{\omega}_x^\varepsilon} - \mathbb{E}[e^{\beta_N \tilde{\omega}_0^\varepsilon}]) f_N(x) \\ & + \frac{1}{\beta_N N^{d/2\alpha}} \mathbb{E}[e^{\beta_N \tilde{\omega}_0^\varepsilon} - 1 - \beta_N \tilde{\omega}_0^\varepsilon] \sum_{\|x\| \leq K\sqrt{N}} f_N(x) + \sum_{\|x\| > K\sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \tilde{\omega}_x} - 1) f_N(x). \end{aligned} \quad (7.36)$$

We control each term separately: the first two terms bring the main contribution; we show that the last three terms can be made arbitrarily small by taking ε small or K large.

1st term in (7.36). We split the term into

$$\sum_{\substack{\varepsilon^\alpha \sqrt{N} \leq \|x\| \leq K\sqrt{N} \\ (\omega_x - \mu) > \varepsilon N^{d/2\alpha}}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \tilde{\omega}_x} - 1) f_N(x) + \sum_{\substack{\|x\| \leq \varepsilon^\alpha \sqrt{N} \\ (\omega_x - \mu) > \varepsilon N^{d/2\alpha}}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \tilde{\omega}_x} - 1) f_N(x). \quad (7.37)$$

For the first term we partition $B_K \times (\varepsilon, \infty)$ into rectangles with diameter $\delta > 0$, where $B_K := \{x \in \mathbb{R}^d : \|x\| \leq K\}$. By denoting this partition \mathcal{P}_δ , π a patch of \mathcal{P}_δ and (x_π, w_π) the center of π , we can write the first term above as

$$\begin{aligned} & \sum_{\pi \in \mathcal{P}_\delta} \sum_{\varepsilon^\alpha \sqrt{N} \leq \|x\| \leq K\sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \tilde{\omega}_x} - 1) f_N(x) \mathbb{1}_{\{(\frac{x}{\sqrt{N}}, \frac{\omega_x}{N^{d/2\alpha}}) \in \pi\}} \\ & = \sum_{\pi \in \mathcal{P}_\delta} (1 + o_\delta(1)) \frac{1}{\beta} (e^{\beta w_\pi} - 1) f(x_\pi) \sum_{\varepsilon^\alpha \sqrt{N} \leq \|x\| \leq K\sqrt{N}} \mathbb{1}_{\{(\frac{x}{\sqrt{N}}, \frac{\omega_x}{N^{d/2\alpha}}) \in \pi\}}, \end{aligned} \quad (7.38)$$

where we also used that $\lim_{N \rightarrow \infty} \beta_N N^{-d/2\alpha} = \beta \in [0, \infty)$ (with by convention $\frac{1}{\beta} (e^{\beta w_\pi} - 1) = w_\pi$ when $\beta = 0$) and $\lim_{N \rightarrow \infty} f_N(z\sqrt{N}) = f(z)$, see (2.4); here $o_\delta(1)$ denotes errors that are negligible as $\delta \searrow 0$ for fixed ε .

It is now not hard to check that (see [19, p. 4036])

$$\sum_{0 < \|x\| \leq K\sqrt{N}} \mathbb{1}_{\{(\frac{x}{\sqrt{N}}, \frac{\omega_x}{N^{d/2\alpha}}) \in \pi\}} \longrightarrow \mathcal{P}(\pi) \quad \text{weakly,} \quad (7.39)$$

where $\mathcal{P}(\pi)$ is a Poisson point measure on π with intensity $\eta(dx, dw) = \alpha w^{-(\alpha+1)} dx dw$.

Thus, by first sending $N \rightarrow \infty$ and then letting $\delta \rightarrow 0$, we have the following convergence in distribution:

$$\sum_{\substack{\varepsilon^\alpha \sqrt{N} \leq \|x\| \leq K\sqrt{N} \\ (\omega_x - \mu) > \varepsilon N^{d/2\alpha}}} \frac{1}{\beta_N N^{d/2\alpha}} \left(e^{\beta_N \tilde{\omega}_x} - 1 \right) f_N(x) \xrightarrow{(d)} \int_{(B_K \setminus B_{\varepsilon^\alpha}) \times (\varepsilon, \infty)} \frac{1}{\beta} (e^{\beta w} - 1) f(x) \mathcal{P}(dx, dw). \quad (7.40)$$

Note that [27, Theorem 10.15] ensures that the integral is well-defined (in particular at $w = +\infty$, recall also the proof of Proposition 2.11 in Section 3.4).

For the second term in (7.37) we show that uniformly on $N \in \mathbb{N}$

$$\mathbb{P} \left(\sum_{\substack{\|x\| \leq \varepsilon^\alpha \sqrt{N} \\ \tilde{\omega}_x > \varepsilon N^{d/2\alpha}}} \frac{1}{\beta_N N^{d/2\alpha}} \left(e^{\beta_N \tilde{\omega}_x} - 1 \right) f_N(x) > \varepsilon^{\alpha/2} \right) \leq c_\alpha \varepsilon^{\alpha/2} \log(1/\varepsilon). \quad (7.41)$$

Since $\tilde{\omega}_x \leq \bar{\omega}_x$, we replace $\tilde{\omega}_x$ by $\bar{\omega}_x$. Moreover, we consider the event $\mathcal{A} = \left\{ \max_{\|x\| \leq \varepsilon^\alpha \sqrt{N}} \bar{\omega}_x \leq N^{d/2\alpha} \right\}$ and split (7.41) according to \mathcal{A} : the probability in the left-hand side of (7.41) is bounded by

$$\mathbb{P} \left(\sum_{\|x\| \leq \varepsilon^\alpha \sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} \left(e^{\beta_N \bar{\omega}_x} - 1 \right) \mathbb{1}_{\{\bar{\omega}_x > \varepsilon N^{d/2\alpha}\}} f_N(x) > \varepsilon^{\alpha/2}; \mathcal{A} \right) + \mathbb{P}(\mathcal{A}^c). \quad (7.42)$$

Since $\beta_N N^{d/2\alpha} \leq c$, we have that on \mathcal{A} , $(e^{\beta_N \bar{\omega}_x} - 1) \leq C \beta_N N^{d/2\alpha}$. Then, note that we have

$$|f_N(x) - f(x/\sqrt{N})| = |v_N \mathbf{P}(x \in \mathcal{R}_N) - f(x/\sqrt{N})| \leq c f(x/\sqrt{N}) \quad (7.43)$$

for all $x \in \mathbb{Z}^d \setminus \{0\}$ with $\|x\| \leq K\sqrt{N}$ by [32, Thm. 1.8] (in dimension $d \geq 3$) and [32, Thm. 1.6] (in dimension $d = 2$) and by a Riemann sum approximation (and dominated convergence), we have

$$\mathbb{E} \left[\sum_{\|x\| \leq \varepsilon^\alpha \sqrt{N}} \mathbb{1}_{\{\bar{\omega}_x > \varepsilon N^{d/2\alpha}\}} f_N(x) \right] \leq C \varepsilon^{-\alpha} N^{-d/2} \sum_{\|x\| \leq \varepsilon^\alpha \sqrt{N}} f_N(x) \leq C' \varepsilon^{-\alpha} \int_{B_{\varepsilon^\alpha}} f(x) dx,$$

where we used that $\mathbb{P}(\bar{\omega}_x > \varepsilon N^{d/2\alpha}) \leq C \varepsilon^{-\alpha} N^{-d/2}$. Then from asymptotic estimates on f , see (3.21), we easily obtain that this is bounded by a constant times $\varepsilon^{2\alpha} \log(1/\varepsilon)$. Therefore, by Markov's inequality, the first term of (7.42) is bounded by $C'' \varepsilon^{-\alpha} \varepsilon^{3\alpha/2} \log(1/\varepsilon) \leq \varepsilon^{\alpha/2} \log(1/\varepsilon)$.

Finally, since $\mathbb{P}(\mathcal{A}^c) = \mathbb{P} \left(\max_{\|x\| \leq \varepsilon^\alpha \sqrt{N}} \bar{\omega}_x > N^{d/2\alpha} \right) \leq c \varepsilon^{\alpha d}$, we complete the proof of (7.41).

2nd term in (7.36). Using that $\omega_0 - \mu$ is centered, we easily get from (1.4) that

$$N^{\frac{d}{2}(1-\frac{1}{\alpha})} \mathbb{E}[\tilde{\omega}_0^\varepsilon] = -N^{\frac{d}{2}(1-\frac{1}{\alpha})} \mathbb{E}[(\omega_0 - \mu) \mathbb{1}_{\{(\omega_0 - \mu) > \varepsilon N^{d/2\alpha}\}}] \xrightarrow{N \rightarrow \infty} -\alpha \int_\varepsilon^\infty w \cdot w^{-(1+\alpha)} dw.$$

Again by (7.43) and a Riemann-sum approximation (and dominated convergence), we get that

$$\lim_{N \rightarrow \infty} N^{-\frac{d}{2}} \sum_{\|x\| \leq K\sqrt{N}} f_N(x) = \int_{\|x\| \leq K} f(x) dx. \quad (7.44)$$

All together, we get that the second term in (7.36) verifies

$$\lim_{N \rightarrow \infty} N^{-d/2\alpha} \mathbb{E}[\tilde{\omega}_0^\varepsilon] \sum_{\|x\| \leq K\sqrt{N}} f_N(x) = - \int_{B_K \times (\varepsilon, +\infty)} w f(x) \eta(dx, dw). \quad (7.45)$$

Combining (7.40) with (7.45), we obtain that the sum of the first term in (7.37) and the second term in (7.36) converges in distribution to

$$\begin{aligned} & \int_{(B_K \setminus B_{\varepsilon\alpha}) \times (\varepsilon, +\infty)} \frac{1}{\beta} (e^{\beta w} - 1 - \beta w) f(x) \mathcal{P}(dx, dw) + \int_{(B_K \setminus B_{\varepsilon\alpha}) \times (\varepsilon, +\infty)} w f(x) (\mathcal{P} - \eta)(dx, dw) \\ & - \int_{B_{\varepsilon\alpha} \times (\varepsilon, \infty)} \omega f(x) \eta(dx, d\omega). \end{aligned} \quad (7.46)$$

Note that the third integral above is bounded by $c_\alpha \varepsilon^{1+\alpha} \log(1/\varepsilon)$. Letting $\varepsilon \downarrow 0$ and $K \uparrow +\infty$ and by (7.41), one recovers \mathcal{W}_β , see Proposition 2.11; when $\beta = 0$, the first term is equal to zero. It therefore remains to control the remaining terms in (7.36).

3rd term in (7.36). We control the second moment of this term, since it has zero mean. By a Taylor expansion, using that $|\beta_N \tilde{\omega}_x^\varepsilon| \leq C\varepsilon\beta$, we have that

$$\mathbb{E} \left[(e^{\beta_N \tilde{\omega}_0^\varepsilon} - \mathbb{E}[e^{\beta_N \tilde{\omega}_0^\varepsilon}])^2 \right] = \mathbb{E}[e^{2\beta_N \tilde{\omega}_0^\varepsilon}] - \mathbb{E}[e^{\beta_N \tilde{\omega}_0^\varepsilon}]^2 \leq C\beta_N^2 \mathbb{E}[(\tilde{\omega}_0^\varepsilon)^2] \leq C'\varepsilon^{2-\alpha} \beta_N^2 N^{\frac{d}{2\alpha}(2-\alpha)},$$

where the last inequality holds since $\alpha < 2$. Hence, recalling that $f_N(x) = v_N \mathbf{P}(x \in \mathcal{R}_N)$, we get that the third term of (7.36) has a second moment

$$\mathbb{E} \left[\left(\sum_{\|x\| \leq K\sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \tilde{\omega}_x^\varepsilon} - \mathbb{E}[e^{\beta_N \tilde{\omega}_x^\varepsilon}]) f_N(x) \right)^2 \right] \leq c\varepsilon^{2-\alpha} N^{-\frac{d}{2}} v_N^2 J_N \leq c'\varepsilon^{2-\alpha}, \quad (7.47)$$

where for the last inequality we used the definition of v_N (see Section 2.3) together with Lemma B.5 for $d = 2, 3$.

4th term in (7.36). Again, by a Taylor expansion, and using $\mathbb{E}[(\tilde{\omega}_0^\varepsilon)^2] \leq c\varepsilon^{2-\alpha} N^{\frac{d}{2\alpha}(2-\alpha)}$, we have

$$\frac{1}{\beta_N N^{d/2\alpha}} \mathbb{E}[e^{\beta_N \tilde{\omega}_0^\varepsilon} - 1 - \beta_N \tilde{\omega}_0^\varepsilon] \leq C \frac{\beta_N}{N^{d/2\alpha}} \mathbb{E}[(\tilde{\omega}_0^\varepsilon)^2] \leq C\varepsilon^{2-\alpha} \beta_N N^{\frac{d}{2\alpha}} N^{-\frac{d}{2}}.$$

Now, using (7.44) and the fact that $\beta_N N^{d/2\alpha} \leq c$, we get that the 4th term in (7.36) verifies

$$0 \leq \frac{1}{\beta_N N^{d/2\alpha}} \mathbb{E}[e^{\beta_N \tilde{\omega}_0^\varepsilon} - 1 - \beta_N \tilde{\omega}_0^\varepsilon] \sum_{\|x\| \leq K\sqrt{N}} f_N(x) \leq C\varepsilon^{2-\alpha}. \quad (7.48)$$

5th term in (7.36). Here, we follow the line of proof of [19, Lem. 5.1]: we prove that (recall the notation $\bar{\omega}_x = \omega_x - \mu$)

$$\mathbb{P} \left(\sum_{\|x\| > K\sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \bar{\omega}_x} - 1) f_N(x) > e^{-K} \right) \leq CK^{d-2\alpha}, \quad (7.49)$$

which can be made arbitrarily small by taking K large, since $\alpha > d/2$. Note that since $\tilde{\omega}_x \leq \bar{\omega}_x$, this enables us to control the last term in (7.36).

To prove (7.49), we decompose the sum according to whether $\bar{\omega}_x$ is smaller or larger than $1/\beta_N$.

To control the part of the sum with $\bar{\omega}_x \leq 1/\beta_N$, we use Lemma C.1 to get that

$$\mathbb{E} \left[\left| e^{\beta_N \bar{\omega}_x \mathbb{1}_{\{\bar{\omega}_x \leq 1/\beta_N\}}} - 1 \right| \right] \leq c' \beta_N^\alpha.$$

Hence, the sum over the terms with $\bar{\omega}_x \leq 1/\beta_N$ has L^1 norm bounded by

$$\begin{aligned} \mathbb{E} \left[\sum_{\|x\| > K\sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} \left| e^{\beta_N \bar{\omega}_x \mathbb{1}_{\{\bar{\omega}_x \leq 1/\beta_N\}}} - 1 \right| f_N(x) \right] &\leq (\beta_N N^{d/2\alpha})^{\alpha-1} N^{-\frac{d}{2}} \sum_{\|x\| > K\sqrt{N}} f_N(x) \\ &\leq c' \int_{\|x\| > K} f(x) dx, \end{aligned}$$

where we used $\beta_N \leq cN^{-d/2\alpha}$ and the same Riemann-sum approximation as in (7.44). Then from asymptotic estimates on f , see (3.20), we easily obtain that this is bounded by a constant times e^{-cK^2} . Hence, applying Markov's inequality, we get that the probability that this part of the sum is larger than $e^{-K}/2$ is bounded by $ce^K e^{-cK^2} \leq cK^{d-2\alpha}$.

Let us now control the sum over the terms with $\bar{\omega}_x > 1/\beta_N$. We write

$$\begin{aligned} &\mathbb{P} \left(\sum_{\|x\| > K\sqrt{N}} \frac{1}{\beta_N N^{d/2\alpha}} (e^{\beta_N \bar{\omega}_x} - 1) \mathbb{1}_{\{\bar{\omega}_x > 1/\beta_N\}} f_N(x) > e^{-K}/2 \right) \\ &\leq \sum_{j=1}^{K^{-1}\sqrt{N}} \mathbb{P} \left(\sum_{\|x\| \in (j, j+1]K\sqrt{N}} e^{\beta_N \bar{\omega}_x} \mathbb{1}_{\{\bar{\omega}_x > 1/\beta_N\}} f_N(x) > \beta_N N^{d/2\alpha} e^{-K} 2^{-j-2} \right). \quad (7.50) \end{aligned}$$

Now, we set $N_j^K = |\{x \in \mathbb{Z}^d, \bar{\omega}_x > 1/\beta_N, \|x\| \in (j, j+1]K\sqrt{N}\}|$, and we bound each probability in the sum by

$$\begin{aligned} &\mathbb{P} \left(N_j^K > e^K j^{1+d} \beta_N N^{d/2\alpha} \right) \\ &+ \mathbb{P} \left(\sum_{\|x\| \in (j, j+1]K\sqrt{N}} e^{\beta_N \bar{\omega}_x} \mathbb{1}_{\{\bar{\omega}_x > 1/\beta_N\}} f_N(x) > \beta_N N^{d/2\alpha} e^{-K} 2^{-j-2}; N_j^K \leq e^K j^{1+d} \beta_N N^{d/2\alpha} \right). \quad (7.51) \end{aligned}$$

First of all, by Markov's inequality, since the number of sites verifying $\|x\| \in (j, j+1]K\sqrt{N}$ is bounded by a constant times $K^d j^{d-1} N^{d/2}$ and since $\mathbb{P}(\bar{\omega}_x > 1/\beta_N) \leq c\beta_N^\alpha$, we get that

$$\mathbb{P} \left(N_j^K > e^K j^{d+1} \beta_N N^{d/2\alpha} \right) \leq ce^{-K} j^{-(d+1)} (\beta_N N^{d/2\alpha})^{-1} \times K^d j^{d-1} N^{\frac{d}{2}} \beta_N^\alpha \leq c' K^d e^{-K} j^{-2}, \quad (7.52)$$

where we used that $\beta_N^{\alpha-1} \leq cN^{d(\alpha-1)/2\alpha}$. For the second probability in (7.51), we use that $f_N(x)$ is bounded above by $e^{-cj^2 K^2}$ uniformly for $\|x\| > jK\sqrt{N}$ (recall (3.20)), we bound the number of terms in the sum by $e^K j^{1+d} \beta_N N^{d/2\alpha}$ (using the condition on N_j^K) and we bound $e^{\beta_N \bar{\omega}_x}$ by its maximum. We then obtain that the second probability in (7.51) is bounded by

$$\mathbb{P} \left(e^K j^{1+d} e^{-cj^2 K^2} \max_{\|x\| \in (j, j+1]K\sqrt{N}} \{e^{\beta_N \bar{\omega}_x}\} > e^{-K} 2^{-j-2} \right).$$

Now, using that $e^{-2K} j^{-(1+d)} 2^{-j} e^{cj^2 K^2} \geq e^{c'j^2 K^2}$ uniformly in $j \geq 1$ and noting that $N^{d/2} \beta_N^\alpha \leq c$, we get by a union bound that the above probability is bounded by

$$c' j^{d-1} K^d N^{\frac{d}{2}} \mathbb{P}(e^{c\beta_N \bar{\omega}_x} > e^{c'j^2 K^2}) \leq c' j^{d-1-2\alpha} K^{d-2\alpha} N^{\frac{d}{2}} \beta_N^\alpha \leq c' j^{d-1-2\alpha} K^{d-2\alpha}. \quad (7.53)$$

All together, summing over j for (7.52) and (7.53) (recall that $\alpha > d/2$), we get that the probability in (7.50) is bounded by a constant times $K^d e^{-K} + K^{d-2\alpha}$. This concludes the proof of (7.49).

Conclusion of the proof. Recall that the sum of the first two terms in (7.36) converge to (7.46). Then, we let $\varepsilon \downarrow 0$ and we use (7.47)-(7.48) to get that the 3rd and 4th term in (7.36) become negligible. Finally, letting $K \uparrow \infty$, (7.49) makes sure that the 5th term in (7.36) goes to 0 in probability.

7.3 Transversal fluctuations

Let us first show that the transversal fluctuations are *at least* of order \sqrt{N} . Since $\bar{Z}_{N,\beta_N}(M_N \leq \eta\sqrt{N}) = \bar{Z}_{N,\beta_N} \bar{\mathbf{P}}_{N,\beta_N}(M_N \leq \eta\sqrt{N})$ and because $\bar{Z}_{N,\beta_N} \rightarrow 1$ in \mathbb{P} -probability, we can focus on $\bar{Z}_{N,\beta_N}(M_N \leq \eta\sqrt{N})$.

By the same argument as (7.6)-(7.7), $\bar{Z}_{N,\beta_N}(M_N \leq \eta\sqrt{N}) - Z_{N,\beta_N}^{\omega}(M_N \leq \eta\sqrt{N}) \rightarrow 0$ in \mathbb{P} -probability. We then have

$$\begin{aligned} \mathbb{P}(\bar{Z}_{N,\beta_N}(M_N \leq \eta\sqrt{N}) > \varepsilon) &\leq \mathbb{P}\left(|\bar{Z}_{N,\beta_N}(M_N \leq \eta\sqrt{N}) - Z_{N,\beta_N}^{\omega}(M_N \leq \eta\sqrt{N})| > \frac{\varepsilon}{2}\right) \\ &\quad + \mathbb{P}\left(Z_{N,\beta_N}^{\omega}(M_N \leq \eta\sqrt{N}) > \frac{\varepsilon}{2}\right), \end{aligned}$$

where the second term is bounded above by $C_\varepsilon(1 - e^{-c\eta^2})$, thanks to the Markov's inequality and reasoning as in (7.5), which can be made arbitrarily small by choosing small enough η : this concludes the proof that transversal fluctuations are at least of order \sqrt{N} .

The proof for transversal fluctuations are *at most* of order \sqrt{N} is very similar, where one just needs to restrict the path to $\{M_N \leq A_N \sqrt{N \log N}\}$ first (see Lemma 7.1 for more details), so we omit it. \square

8 Region C: proof of Theorem 2.15

Here we have $\lim_{N \rightarrow \infty} \beta_N N^{\frac{d}{\alpha}-1} = \beta \in [0, +\infty)$, and we work conditionally on the event $\beta \leq \beta_c$, i.e. $\hat{\mathcal{T}}_\beta = 0$ in Theorem 2.2. Recall that if $\alpha > 1$ we set $\mu = \mathbb{E}[\omega_0]$, whereas when $\alpha < 1$ we let μ be any real number; we again use the notation $\bar{Z}_{N,\beta_N} = Z_{N,\beta_N}^{\omega, h=\mu}$.

8.1 Preliminary: paths cannot stay at scale N nor at an intermediate scale

First of all, let us stress that conditionally on having $\hat{\mathcal{T}}_\beta = 0$, paths cannot stay at scale N . Indeed, similarly to what is done in Section 5 (see in particular (5.7)), we have that for any $\eta > 0$

$$\frac{1}{\beta_N N^{d/\alpha}} \log Z_{N,\beta_N}^{\omega, h} (M_N \geq \eta N) \longrightarrow \hat{\mathcal{T}}_{\beta, \geq \eta} := \sup_{s \in \mathcal{D}, \sup_{t \in [0,1]} \|s(t)\| \geq \eta, \hat{\mathbf{E}}_{\text{ent}}(s) < +\infty} \left\{ \pi(s) - \frac{1}{\beta} \hat{\mathbf{E}}_{\text{ent}}(s) \right\}.$$

Now, on the event that $\hat{\mathcal{T}}_\beta = 0$ we have $\hat{\mathcal{T}}_{\beta, \geq \eta} < 0$ a.s., see Lemma A.5 in Appendix. Hence, if $\lim_{N \rightarrow \infty} \beta_N N^{d/\alpha-1} = \beta > 0$, we get that for any $\eta > 0$,

$$\lim_{N \rightarrow \infty} \mathbb{P}\left(Z_{N,\beta_N}^{\omega, h} (M_N \geq \eta N) \geq e^{-N^{1/2}} \mid \hat{\mathcal{T}}_\beta = 0\right) = 0. \quad (8.1)$$

If $\lim_{N \rightarrow \infty} \beta_N N^{d/\alpha-1} = 0$, in view of Remark 2.6, one has

$$\lim_{N \rightarrow \infty} \frac{1}{N} \log Z_{N,\beta_N}^{\omega, h} (M_N \geq \eta N) = \lim_{N \rightarrow \infty} \frac{1}{N} \log \mathbf{P}(M_N \geq \eta N) < 0,$$

so the conclusion 8.1 remains valid.

Let us now prove here the analogous of Proposition 6.1, stating that paths cannot stay at some intermediate scale between $N^{1/2}$ and N .

Lemma 8.1. Assume that $\lim_{N \rightarrow \infty} \beta_N N^{\frac{d-\alpha}{2}} = \beta \in [0, +\infty)$. Then, there exists some $\nu' > 0$ and a constant c_1 such that for any $A_N \in [\sqrt{\log N}, N^{1/4}]$ and any $\eta \in (0, 1)$ we have that for N sufficiently large

$$\mathbb{P}\left(\log \bar{Z}_{N,\beta_N}(M_N \in [A_N \sqrt{N}, \eta N]) > -c_1 A_N^2\right) \leq c\eta^{\nu'} + cA_N^{-\nu'}.$$

Proof. The proof is very similar to that of Proposition 6.1. First, we divide the partition function as

$$\bar{Z}_{N,\beta_N}(M_N \in [A_N \sqrt{N}, \eta N]) = \bar{Z}_{N,\beta_N}(M_N \in [A_N \sqrt{N}, N^{3/4}]) + \bar{Z}_{N,\beta_N}(M_N \in [N^{3/4}, \eta N]). \quad (8.2)$$

For the first term, we divide the partition function as

$$\begin{aligned} \bar{Z}_{N,\beta_N}(M_N \in [A_N \sqrt{N}, N^{3/4}]) &= \sum_{k=\log_2 A_N}^{\frac{1}{4} \log_2 N} \bar{Z}_{N,\beta_N}(M_N \in [2^k \sqrt{N}, 2^{k+1} \sqrt{N}]) \\ &\leq \sum_{k=\log_2 A_N}^{\frac{1}{4} \log_2 N} e^{-c2^{2k}} \sqrt{\bar{Z}_{N,\beta_N}(M_N \leq 2^{k+1} \sqrt{N})}. \end{aligned} \quad (8.3)$$

where we used Cauchy–Schwarz inequality for the second inequality, together with the fact that $\mathbb{P}(M_N \geq 2^k \sqrt{N}) \leq \exp(-c2^{2k})$. Now, applying Lemma 6.2 with $h_N = 2^k \sqrt{N}$, we get for the range of k considered that for any fixed constant $c_0 > 0$

$$\mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \leq 2^{k+1} \sqrt{N}) \geq \exp(c_0 2^{2k})\right) \leq c'' N^{-\frac{1}{4}(\frac{d}{\alpha}-2)\nu} + c' 2^{-2\nu k},$$

where we used that $\varepsilon_N = N\beta_N h_N^{d/\alpha-2} \leq c_\beta N^{-\frac{1}{4}(d/\alpha-2)}$ since $h_N \leq N^{3/4}$ and $\alpha < \frac{d}{2}$.

Together with (8.3), a union bound therefore shows that, provided that c_1 has been chosen small enough, we have

$$\mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \in [A_N \sqrt{N}, N^{3/4}]) \geq e^{-c_1 A_N^2}\right) \leq c'' A_N^{-2\nu} + c(\log N) N^{-\frac{1}{4}(\frac{d}{\alpha}-2)}. \quad (8.4)$$

Note that the second term is bounded by $A_N^{-\nu'}$ since $A_N \leq N^{1/4}$.

For the second term in (8.2), we proceed analogously: we decompose it as

$$\begin{aligned} \bar{Z}_{N,\beta_N}(M_N \in [N^{3/4}, \eta N]) &= \sum_{k=-\log_2 \eta}^{\frac{1}{4} \log_2 N} \bar{Z}_{N,\beta_N}(M_N \in [2^{-k-1} N, 2^{-k} N]) \\ &\leq \sum_{k=-\log_2 \eta}^{\frac{1}{4} \log_2 N} e^{-c2^{-2k} N} \bar{Z}_{N,\beta_N}(M_N \leq 2^{-k} N). \end{aligned} \quad (8.5)$$

Again, we can apply Lemma 6.2 with $h_N = 2^{-k} N$ to get for the range of k considered that

$$\mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \leq 2^{-k} N) \geq \exp(c_0 2^{-2k} N)\right) \leq c2^{-k(\frac{d}{\alpha}-2)\nu} + N^{-\nu/2}, \quad (8.6)$$

where we used $\varepsilon_N = N\beta_N h_N^{d/\alpha-2} \leq c_\beta 2^{-k(d/\alpha-2)}$, together with the fact that $h_N^2/N \geq N^{1/2}$.

Again, together with (8.5), a union bound therefore shows that, provided that c_1 has been chosen small enough, we have

$$\mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \in [N^{3/4}, \eta N]) \geq e^{-c_1 N^{1/2}}\right) \leq c'' \eta^{(\frac{d}{\alpha}-2)\nu} + c(\log N)N^{-\nu/2}. \quad (8.7)$$

Note again that the second term is bounded by $A_N^{-\nu'}$ since $A_N \leq N^{1/4}$. Combining (8.4) and (8.7) therefore concludes the proof of Lemma 8.1. \square

8.2 The case $\alpha \in (0, 1)$

We first prove (2.11), in the case $\alpha \in (0, 1)$. We show the following: if $\alpha \in (0, 1)$, then for any $h \in \mathbb{R}$,

$$\frac{v_N}{\beta_N N^{d/2\alpha}} \left(Z_{N,\beta_N}^{\omega,h} - 1 \right) \xrightarrow{(d)} \mathcal{W}_0. \quad (8.8)$$

Recall that $v_N = \log N$ in dimension $d = 2$ and $v_N = N^{\frac{d}{2}-1}$ in dimension $d \geq 3$.

Proof of (8.8). First of all, using (8.1) and Lemma 8.1, we get that for any $a > 0$ there is some $A > 0$ such that

$$\lim_{N \rightarrow \infty} \mathbb{P}\left(N^a \left(Z_{N,\beta_N}^{\omega,h} - Z_{N,\beta_N}^{\omega,h}(M_N \leq A\sqrt{N \log N}) \right) > \varepsilon \mid \widehat{\mathcal{T}}_\beta = 0\right) = 0.$$

Therefore, we may focus on the convergence of $Z_{N,\beta_N}^{\omega,h}(M_N \leq A\sqrt{N \log N})$. We have

$$e^{-|h|\beta_N N} \leq \frac{Z_{N,\beta_N}^{\omega,h}(M_N \leq A\sqrt{N \log N})}{Z_{N,\beta_N}^{\omega,h=0}(M_N \leq A\sqrt{N \log N})} \leq e^{|h|\beta_N N},$$

so that, for any $h \in \mathbb{R}$, noting that $\beta_N N \leq cN^{2-d/\alpha} \rightarrow 0$, we have

$$\left| \frac{Z_{N,\beta_N}^{\omega,h}(M_N \leq A\sqrt{N \log N})}{Z_{N,\beta_N}^{\omega,h=0}(M_N \leq A\sqrt{N \log N})} - 1 \right| \leq C|h|\beta_N N. \quad (8.9)$$

Notice that $v_N/(\beta_N N^{d/2\alpha}) \times \beta_N N \leq N^{\frac{d}{2}(1-\frac{1}{\alpha})} \log N$ and goes to 0 as $N \rightarrow \infty$ (the factor $\log N$ factor is present only in dimension $d = 2$). Hence, in view of (8.9), we therefore only have to prove the result for $h = 0$, that is

$$\frac{v_N}{\beta_N N^{d/2\alpha}} \left(Z_{N,\beta_N}^{\omega,h=0}(M_N \leq A\sqrt{N \log N}) - 1 \right) \xrightarrow{(d)} \mathcal{W}_0, \quad (8.10)$$

since it also implies that $Z_{N,\beta_N}^{\omega,h=0}(M_N \leq A\sqrt{N \log N})$ goes to 1 in probability.

Now, since the ω 's are non-negative, on the event $M_N \leq A\sqrt{N \log N}$ we have

$$\begin{aligned} \mathbb{P}\left(\sum_{x \in \mathcal{R}_N} \omega_x > N^{\frac{d}{2\alpha}} (\log N)^{\frac{d}{\alpha}}\right) &\leq \mathbb{P}\left(\sum_{\|x\| \leq A\sqrt{N \log N}} \omega_x > N^{\frac{d}{2\alpha}} (\log N)^{\frac{d}{\alpha}}\right) \\ &\leq (A\sqrt{N \log N})^d \mathbb{P}(\omega_0 > N^{\frac{d}{2\alpha}} (\log N)^{\frac{d}{\alpha}}) \leq C A^d (\log N)^{-\frac{d}{2}}, \end{aligned}$$

where we used a union bound for the second inequality. Hence, with \mathbb{P} probability going to 1, we have that $\beta_N \sum_{x \in \mathcal{R}_N} \omega_x \leq \beta_N N^{\frac{d}{2\alpha}} (\log N)^{\frac{d}{\alpha}} \rightarrow 0$ on the event $\mathcal{A}_N := \{M_N \leq A\sqrt{N \log N}\}$. By a Taylor expansion, we therefore have

$$\left| Z_{N,\beta_N}^{\omega,h=0}(M_N \leq A\sqrt{N \log N}) - \mathbb{P}(\mathcal{A}_N) - \beta_N \mathbb{E}\left[\sum_{x \in \mathcal{R}_N} \omega_x \mathbb{1}_{\{M_N \leq A\sqrt{N \log N}\}}\right] \right| \leq \left(\beta_N N^{\frac{d}{2\alpha}} (\log N)^{\frac{d}{\alpha}} \right)^2,$$

with \mathbb{P} -probability going to 1 as $N \rightarrow \infty$. Notice that since $\beta_N \leq cN^{1-d/\alpha}$ and $v_N \leq N^{d/2-1} \log N$, we get that

$$\frac{v_N}{\beta_N N^{d/2\alpha}} \times (\beta_N N^{\frac{d}{2\alpha}} (\log N)^{\frac{d}{\alpha}})^2 \leq N^{\frac{d}{2}(1-\frac{1}{\alpha})} (\log N)^{\frac{2d}{\alpha}},$$

which goes to 0 as $N \rightarrow \infty$ since $\alpha < 1$.

Note that $1 - \mathbf{P}(\mathcal{A}_N) \leq N^{-cA^2}$ and A is large enough. All together, it boils down to proving that

$$\frac{v_N}{N^{d/2\alpha}} \mathbf{E} \left[\sum_{x \in \mathcal{R}_N} \omega_x \mathbb{1}_{\{M_N \leq A\sqrt{N \log N}\}} \right] \xrightarrow{(d)} \mathcal{W}_0. \quad (8.11)$$

Recalling the notation $f_N(x) = v_N \mathbf{P}(x \in \mathcal{R}_N)$, we have the following bound

$$\begin{aligned} & \left| \frac{v_N}{N^{d/2\alpha}} \sum_{x \in \mathcal{R}_N} \omega_x \mathbf{P}(x \in \mathcal{R}_N, M_N \leq A\sqrt{N \log N}) - \sum_{\|x\| \leq A\sqrt{N \log N}} \frac{\omega_x}{N^{d/2\alpha}} f_N(x) \right| \\ & \leq \frac{v_N}{N^{d/2\alpha}} \max_{\|x\| \leq A\sqrt{N \log N}} \{\omega_x\} \mathbf{E} [|\mathcal{R}_N| \mathbb{1}_{\{M_N > A\sqrt{N \log N}\}}] \\ & \leq N^{\frac{d}{2}} \max_{\|x\| \leq A\sqrt{N \log N}} \left\{ \frac{\omega_x}{N^{d/2\alpha}} \right\} \left(\mathbf{P}(M_N > A\sqrt{N \log N}) + \log N \mathbf{P}(|\mathcal{R}_N| > AN/\log N) \mathbb{1}_{d=2} \right), \end{aligned}$$

which goes to 0 in probability provided that A has been fixed large enough. Therefore, the convergence (8.11) is equivalent to showing

$$\sum_{\|x\| \leq A\sqrt{N \log N}} \frac{\omega_x}{N^{d/2\alpha}} f_N(x) \xrightarrow{(d)} \mathcal{W}_0 = \int_{\mathbb{R}^d \times \mathbb{R}^+} w f(x) \mathcal{P}(dx, dw), \quad (8.12)$$

with \mathcal{P} a Poisson point process on $\mathbb{R}^d \times \mathbb{R}^+$ with intensity $\alpha w^{-\alpha-1} dx dw$. The proof is very similar to the proof of Lemma 7.6 — see in particular the method used in (7.38)–(7.39).

Since $w^{-\alpha}$ is integrable around 0 when $\alpha \in (0, 1)$, by the same argument as in (7.38)–(7.39), we have analogously to (7.40)

$$\sum_{\|x\| \leq K\sqrt{N}} \frac{\omega_x}{N^{d/2\alpha}} f_N(x) \xrightarrow{(d)} \int_{B_K \times \mathbb{R}^+} w f(x) \mathcal{P}(dx, dw). \quad (8.13)$$

For the remaining term in (8.12), we get that it verifies

$$\mathbb{P} \left(\sum_{K\sqrt{N} < \|x\| \leq A\sqrt{N \log N}} \frac{\omega_x}{N^{d/2\alpha}} f_N(x) > K^{-1} \right) \leq \sum_{k=1}^{\infty} \mathbb{P} \left(\sum_{\|x\| \in (2^{k-1}, 2^k] K\sqrt{N}} \omega_x f_N(x) > 2^{-k} K^{-1} N^{\frac{d}{2\alpha}} \right). \quad (8.14)$$

Since $f_N(x) \leq C e^{-c\|x\|^2/N}$, we get that this probability is bounded by a constant times

$$\sum_{k=1}^{\infty} 2^{kd} K^d N^{\frac{d}{2}} \mathbb{P} \left(\omega_0 > c' 2^{-k} K^{-1} e^{c2^{2(k-1)} K^2} N^{\frac{d}{2\alpha}} \right) \leq c \sum_{k=1}^{\infty} (2^k K)^{d+\alpha} e^{-c_\alpha 2^k K^2} \leq c e^{-c' K^2}. \quad (8.15)$$

Combining this with (8.13) we obtain (8.12) by letting $N \rightarrow \infty$ and then $K \rightarrow \infty$. \square

8.3 The case $\alpha \in (1, \frac{d}{2})$

We now prove (2.10) and (2.11) in the case $\alpha \in (1, \frac{d}{2})$; in particular $d \geq 3$. The proof is similar to what is done in Section 7: first we truncate the environment and then we prove the convergence of the partition function with truncated environment.

Step 1. Truncation of the environment. Let us define the truncation level

$$k_N := (\log N)^{\frac{d}{\alpha}} N^{\frac{d}{2\alpha}}, \quad (8.16)$$

and the truncated environment $\tilde{\omega} = (\omega - \mu)\mathbb{1}_{\{\omega \leq k_N\}}$. Let us stress that $\beta_N k_N \leq c(\log N)^{d/\alpha} N^{1-d/2\alpha}$ goes to 0 as $N \rightarrow \infty$, since $\alpha < d/2$. We again define $\lambda_N := \log \mathbb{E}[e^{\beta_N \tilde{\omega}}]$ and we notice that thanks to Lemma C.1 (in the case $\beta_N k_N < 1$) we have

$$\lambda_N \leq C\beta_N k_N^{1-\alpha} \leq CN^{1-\frac{d}{2}-\frac{d}{2\alpha}}(\log N)^C. \quad (8.17)$$

In particular, $\lambda_N N$ goes to 0 as $N \rightarrow \infty$ since $\alpha < d/2$.

Lemma 8.2. *Let $\alpha \in (1, \frac{d}{2})$. Assume that $\lim_{N \rightarrow \infty} \beta_N N^{\frac{d-\alpha}{\alpha}} = \beta \in [0, +\infty)$. Then for any $a > 0$, conditionally on $\hat{\mathcal{T}}_\beta = 0$, $N^a(\bar{Z}_{N,\beta_N} - Z_{N,\beta_N}^{\tilde{\omega}})$ goes to 0 in probability.*

Proof. Using (8.1) and Lemma 8.1, we get that for any $a > 0$ there is some $A > 0$ such that

$$\lim_{N \rightarrow \infty} \mathbb{P}\left(N^a(\bar{Z}_{N,\beta_N} - \bar{Z}_{N,\beta_N}(M_N \leq A\sqrt{N \log N})) > \varepsilon \mid \hat{\mathcal{T}}_\beta = 0\right) = 0.$$

Now, using the definition (8.16), we have that

$$\mathbb{P}\left(\exists x, \|x\| \leq A\sqrt{N \log N}, \tilde{\omega}_x \neq \omega_x - \mu\right) \leq A^d(N \log N)^{d/2} \mathbb{P}(\omega > k_N) \leq c'(\log N)^{-\frac{d}{2}}$$

Therefore, we get that

$$\lim_{N \rightarrow \infty} \mathbb{P}\left(\bar{Z}_{N,\beta_N}(M_N \leq A\sqrt{N \log N}) \neq Z_{N,\beta_N}^{\tilde{\omega}}(M_N \leq A\sqrt{N \log N})\right) = 0.$$

It therefore only remains to observe that, by Markov's inequality,

$$\begin{aligned} \mathbb{P}\left(N^a Z_{N,\beta_N}^{\tilde{\omega}}(M_N > A\sqrt{N \log N}) > \varepsilon\right) &\leq \varepsilon^{-1} N^a \mathbb{E}[Z_{N,\beta_N}^{\tilde{\omega}}(M_N > A\sqrt{N \log N})] \\ &\leq \varepsilon^{-1} e^{\lambda_N N} N^a \mathbf{P}(M_N > A\sqrt{N \log N}). \end{aligned}$$

Hence, we get that

$$\mathbb{P}\left(N^a Z_{N,\beta_N}^{\tilde{\omega}}(M_N > A\sqrt{N \log N}) > \varepsilon\right) \leq C\varepsilon^{-1} N^a \mathbf{P}(M_N > A\sqrt{N \log N}),$$

which goes to zero provided A is fixed large enough. This concludes the proof. \square

Step 2. Convergence of the truncated partition function. Thanks to Lemma 8.2, it therefore only remains to prove the convergence of the truncated partition function.

Let us stress that when $d \geq 5$ and $\alpha \in (\frac{d}{d-2}, \frac{d}{2})$, then we may apply Lemma 7.2: the only requirement is to have $\beta_N \leq cN^{-d/2\alpha}$ and $\beta_N k_N \leq c(\log N)^\eta$ for some $\eta < 1$ (to apply the bounds (7.9)-(7.10)), which is clearly the case here. This proves the convergence (2.10) in Theorem 2.15.

We therefore focus on the case $\alpha \in (1, \frac{d}{2} \wedge \frac{d}{d-2})$: to conclude the proof of the convergence (2.11) in Theorem 2.15, we need to prove the following:

$$\frac{v_N}{\beta_N N^{d/2\alpha}} (Z_{N,\beta_N}^{\tilde{\omega}} - 1) \xrightarrow{(d)} \mathcal{W}_0, \quad (8.18)$$

with \mathcal{W}_0 defined in (2.9). This is analogous to Lemma 7.5, and let us stress that this lemma uses the condition $\alpha > d/2$ only to control (7.49). Let us rapidly go over the different steps of the proof, the main changes occurring in the proof of Lemma 8.3 below, which is the analogous of Lemma 7.6.

Recall that we define $\lambda_N := \log \mathbb{E}[e^{\beta_N \tilde{\omega}}]$ and $\varepsilon_x := e^{\beta_N \tilde{\omega}_x - \lambda_N} - 1$; let us stress that using that $\beta_N k_N$ goes to 0, a simple Taylor expansion gives that $\mathbb{E}[\varepsilon_0^2] \leq C\beta_N^2 \mathbb{E}[\tilde{\omega}_x^2] + C\lambda_N^2 \leq C'\beta_N^2 k_N^{2-\alpha}$.

We perform the same chaos expansion as (7.26) to get

$$Z_{N,\beta_N}^{\tilde{\omega}} - 1 = \mathbf{E}[e^{\lambda_N |\mathcal{R}_N|} - 1] + \sum_{x \in \mathbb{Z}^d} \varepsilon_x \mathbf{E}[e^{\lambda_N |\mathcal{R}_N|} \mathbb{1}_{\{x \in \mathcal{R}_N\}}] + \mathbf{Y}_N, \quad (8.19)$$

where \mathbf{Y}_N is defined as (7.12).

Note that \mathbf{Y}_N is centered, and following the same computation as in (7.27)-(7.14), we have that $\mathbb{E}[\mathbf{Y}_N^2] \leq C(\mathbb{E}[(\varepsilon_0)^2]J_N)^2 \leq C\beta_N^4 k_N^{4-2\alpha} J_N^2$. Hence, in view of the definition of $k_N = (\log N)^{d/\alpha} N^{d/2\alpha}$, of the definition of $v_N = N^{d/2-1}$ and $J_N \leq cN^{1/2}$ (recall $d \geq 3$), we get that

$$\left(\frac{v_N}{\beta_N N^{d/2\alpha}}\right)^2 \mathbb{E}[\mathbf{Y}_N^2] \leq C\beta_N^2 N^{\frac{d}{\alpha}-\frac{3}{2}} (\log N)^C \leq C' N^{\frac{1}{2}-\frac{d}{\alpha}} (\log N)^C, \quad (8.20)$$

which goes to 0 as $N \rightarrow \infty$ since $\alpha < d$. Therefore, $v_N \mathbf{Y}_N / \beta_N N^{d/2\alpha}$ goes to 0 in probability.

We now handle the first two terms on the right-hand side of (8.19): we use the same decomposition as in (7.31) and the same notation **I**, **II**, **III** for the three terms. We now control the different terms.

For the term **I**, using (8.17) (in particular $\lambda_N N \rightarrow 0$) we have that $\mathbf{I} \leq c(\lambda_N N)^2 \leq N^2 \beta_N^2 k_N^{2-2g\alpha}$. Using the definition of v_N and k_N and since $\beta_N \leq cN^{1-d/\alpha}$, we therefore get that

$$\frac{v_N}{\beta_N N^{d/2\alpha}} \mathbf{I} \leq CN^{2-\frac{d}{2\alpha}-\frac{d}{2}} (\log N)^C \leq C' N^{\frac{1}{2}-\frac{d}{\alpha}} (\log N)^C, \quad (8.21)$$

which goes to 0 since $\alpha < d$.

For the term **II**, it is centered and we have $\mathbb{E}[(\mathbf{II})^2] \leq C\mathbb{E}[(\varepsilon_0)^2](\lambda_N N)^2 J_N \leq C'\beta_N^4 k_N^{4-3\alpha} N^2 J_N$. Using the definition of v_N and k_N , we therefore get that

$$\left(\frac{v_N}{\beta_N N^{d/2\alpha}}\right)^2 \mathbb{E}[(\mathbf{II})^2] \leq C\beta_N^2 N^{\frac{d}{\alpha}-\frac{d}{2}} J_N (\log N)^C \leq N^{2-\frac{d}{2}-\frac{d}{\alpha}} (\log N)^C \leq N^{2-\frac{d}{2}-(2\vee d-2)+\varepsilon}, \quad (8.22)$$

where we used that $J_N \leq cN^{1/2}$ and $\alpha < \frac{d}{2} \wedge \frac{d}{d-2}$. This goes to 0 as $N \rightarrow \infty$ since $d \geq 3$.

For the term **III**, it is easy to see that since $\beta_N \tilde{\omega}_0 \leq \beta_N k_N \rightarrow 0$, we have by a Taylor expansion

$$\left| e^{\lambda_N \mathbf{III}} - \beta_N \sum_{x \in \mathbb{Z}^d} \tilde{\omega}_x \mathbf{P}(x \in \mathcal{R}_N) \right| \leq c\beta_N^2 \sum_{x \in \mathbb{Z}^d} \tilde{\omega}_x^2 \mathbf{P}(x \in \mathcal{R}_N).$$

Hence, using that $\mathbb{E}[\omega_x^2] \leq k^{2-\alpha}$ we obtain that

$$\frac{v_N}{\beta_N N^{d/2\alpha}} \mathbb{E} \left[\left| e^{\lambda_N \mathbf{III}} - \beta_N \sum_{x \in \mathbb{Z}^d} \tilde{\omega}_x \mathbf{P}(x \in \mathcal{R}_N) \right| \right] \leq c \frac{v_N}{N^{d/2\alpha}} \beta_N k_N^{2-\alpha} N \leq cN^{1-\frac{d}{2\alpha}} (\log N)^C, \quad (8.23)$$

which goes to 0 as $N \rightarrow \infty$ since $\alpha < d/2$.

All together, the proof of (8.18) boils down to showing the following lemma

Lemma 8.3. *If $\alpha \in (1, \frac{d}{2})$ and $d < \frac{2\alpha}{\alpha-1}$ (or equivalently $\alpha \in (1, \frac{d}{2} \wedge \frac{d}{d-2})$), we have the convergence*

$$\frac{v_N}{N^{d/2\alpha}} \sum_{x \in \mathbb{Z}^d} \tilde{\omega}_x \mathbf{P}(x \in \mathcal{R}_N) \xrightarrow{(d)} \mathcal{W}_0. \quad (8.24)$$

Proof. The following procedure is similar to the proof of Lemma 7.6. Note that Lemma 7.6 uses the condition $\alpha > d/2$, in particular to ensure that the limit \mathcal{W}_β is finite. Here, the condition $\alpha > d/(d-2)$ also appears naturally, in particular to ensure that the limit \mathcal{W}_0 is finite, see Proposition 2.14.

We write $f_N(x) = v_N \mathbf{P}(x \in \mathcal{R}_N)$ and we use the following decomposition: for a fixed constant K (large)

$$\frac{v_N}{N^{d/2\alpha}} \sum_{x \in \mathbb{Z}^d} \tilde{\omega}_x \mathbf{P}(x \in \mathcal{R}_N) = \frac{1}{N^{d/2\alpha}} \sum_{\|x\| \leq K\sqrt{N}} \tilde{\omega}_x f_N(x) + \frac{1}{N^{d/2\alpha}} \sum_{\|x\| > K\sqrt{N}} \tilde{\omega}_x f_N(x).$$

For the second term, we proceed as in (8.14)-(8.15) to get that it is bounded by K^{-1} with probability larger than $1 - ce^{-c'K^2}$, uniformly in N .

For the first term, then defining $\tilde{\omega}_x^\varepsilon := (\omega_x - \mu) \mathbb{1}_{\{\omega_x \leq \varepsilon N^{d/2\alpha}\}}$, we decompose it further as

$$\frac{1}{N^{d/2\alpha}} \sum_{\substack{\|x\| \leq K\sqrt{N} \\ (\omega_x - \mu) > \varepsilon N^{d/2\alpha}}} \tilde{\omega}_x f_N(x) + \frac{1}{N^{d/2\alpha}} \sum_{\|x\| \leq K\sqrt{N}} \mathbb{E}[\tilde{\omega}_x^\varepsilon] f_N(x) + \frac{1}{N^{d/2\alpha}} \sum_{\|x\| \leq K\sqrt{N}} (\tilde{\omega}_x^\varepsilon - \mathbb{E}[\tilde{\omega}_x^\varepsilon]) f_N(x). \quad (8.25)$$

By the same argument as for (7.40) and (7.45) in Lemma 7.6, we get that the sum of the first two terms converge in distribution to

$$\int_{B_K \times (\varepsilon, \infty)} wf(x)(\mathcal{P} - \eta)(dx, dw). \quad (8.26)$$

Note that it is then valid to send $\varepsilon \downarrow 0$ and $K \uparrow \infty$ in (8.26), since $\alpha < d/(d-2)$ (recall Proposition 2.14). It therefore remains to prove that the last term in (8.25) is negligible.

To conclude the proof, we will show that for any $\varepsilon > 0$, there exists M_ε which verifies $\lim_{\varepsilon \downarrow 0} M_\varepsilon = +\infty$, such that

$$\mathbb{P}\left(\sum_{\|x\| \leq K\sqrt{N}} N^{-\frac{d}{2\alpha}} \bar{\omega}_x^\varepsilon f_N(x) \geq M_\varepsilon^{-\frac{1}{2}(\alpha-1)}\right) \leq CM_\varepsilon^{-\frac{1}{2}(\alpha-1)}, \quad (8.27)$$

where we denoted $\bar{\omega}_x^\varepsilon = \tilde{\omega}_x^\varepsilon - \mathbb{E}[\tilde{\omega}_x^\varepsilon]$ for simplicity. Then, sending $N \rightarrow \infty$ and then letting $\varepsilon \downarrow 0$ and $K \uparrow \infty$, this concludes the proof of Lemma 8.3.

To prove (8.27), we split the sum according to whether $\bar{\omega}_x^\varepsilon f_N(x)$ is larger or smaller than $M_\varepsilon N^{d/2\alpha}$. For the first sum, we use Markov's inequality to get that

$$\begin{aligned} \mathbb{P}\left(\sum_{\|x\| \leq K\sqrt{N}} N^{-\frac{d}{2\alpha}} \bar{\omega}_x^\varepsilon f_N(x) \mathbb{1}_{\{\bar{\omega}_x^\varepsilon f_N(x) > M_\varepsilon N^{d/2\alpha}\}} \geq \frac{1}{2} M_\varepsilon^{-\frac{1}{2}(\alpha-1)}\right) \\ \leq M_\varepsilon^{\frac{1}{2}(\alpha-1)} \sum_{\|x\| \leq K\sqrt{N}} f_N(x) N^{-\frac{d}{2\alpha}} \mathbb{E}\left[\bar{\omega}_x^\varepsilon \mathbb{1}_{\{\bar{\omega}_x^\varepsilon > M_\varepsilon N^{d/2\alpha} f_N(x)^{-1}\}}\right]. \end{aligned} \quad (8.28)$$

Notice that $f_N(x) \leq v_N = N^{d/2-1}$, so $N^{d/2\alpha} f_N(x)^{-1} \geq N^{1-\frac{d}{2\alpha}(\alpha-1)}$ and goes to $+\infty$ as $N \rightarrow \infty$ since we have $d < 2\alpha/(\alpha-1)$: hence we can use the following estimate

$$\mathbb{E}\left[\bar{\omega}_x^\varepsilon \mathbb{1}_{\{\bar{\omega}_x^\varepsilon > M_\varepsilon N^{d/2\alpha} f_N(x)^{-1}\}}\right] \leq CM_\varepsilon^{1-\alpha} N^{\frac{d}{2\alpha}(1-\alpha)} f_N(x)^{\alpha-1}. \quad (8.29)$$

This shows that (8.28) is bounded by

$$CM_\varepsilon^{-\frac{1}{2}(\alpha-1)} N^{-\frac{d}{2}} \sum_{\|x\| \leq K\sqrt{N}} f_N(x)^\alpha \leq CM_\varepsilon^{-\frac{1}{2}(\alpha-1)},$$

where we have used an integral comparison, using that $\int_{\mathbb{R}^d} f(x)^\alpha dx < \infty$ since $\alpha(d-2) < d$, recall (3.21)(3.20).

It remains to control the sum with terms $\bar{\omega}^\varepsilon f_N(x)$ smaller than $M_\varepsilon N^{d/2\alpha}$. With an exponential Chernov's bound (and using the independence of the different terms in the sum), we get

$$\begin{aligned} \mathbb{P}\left(\sum_{\|x\| \leq K\sqrt{N}} N^{-\frac{d}{2\alpha}} \bar{\omega}_x^\varepsilon f_N(x) \mathbb{1}_{\{\bar{\omega}_x^\varepsilon f_N(x) \leq M_\varepsilon N^{d/2\alpha}\}} \geq M_\varepsilon^{-\frac{1}{2}(\alpha-1)}\right) \\ \leq \exp\left(-M_\varepsilon^{\frac{1}{2}(\alpha-1)}\right) \prod_{\|x\| \leq K\sqrt{N}} \mathbb{E}\left[\exp\left(M_\varepsilon^{\alpha-1} N^{-\frac{d}{2\alpha}} \bar{\omega}_x^\varepsilon f_N(x) \mathbb{1}_{\{\bar{\omega}_x^\varepsilon f_N(x) \leq M_\varepsilon N^{d/2\alpha}\}}\right)\right]. \end{aligned} \quad (8.30)$$

We now estimate the exponential moment. Let $\zeta \in (1, 2)$ be such that $\alpha < \zeta < \frac{d}{d-2}$. We write that

$$\begin{aligned} \mathbb{E}\left[\exp\left(M_\varepsilon^{\alpha-1} N^{-\frac{d}{2\alpha}} \bar{\omega}_x^\varepsilon f_N(x) \mathbb{1}_{\{\bar{\omega}_x^\varepsilon f_N(x) \leq M_\varepsilon N^{d/2\alpha}\}}\right)\right] \\ \leq 1 + M_\varepsilon^{\alpha-1} N^{-\frac{d}{2\alpha}} f_N(x) \mathbb{E}\left[\bar{\omega}_x^\varepsilon \mathbb{1}_{\{\bar{\omega}_x^\varepsilon f_N(x) \leq M_\varepsilon N^{d/2\alpha}\}}\right] + CM_\varepsilon^{\zeta(\alpha-1)} N^{-\frac{d}{2\alpha}\zeta} \mathbb{E}\left[|\bar{\omega}_x^\varepsilon|^\zeta\right] f_N(x)^\zeta, \\ \leq 1 + C'M_\varepsilon^{2(\alpha-1)} \varepsilon^{\zeta-\alpha} N^{-\frac{d}{2}} f_N(x)^\zeta, \end{aligned}$$

where the first term is negative for large enough N , since $\mathbb{E}[\bar{\omega}_x^\varepsilon] = 0$, $\mathbb{E}[\bar{\omega}_x^\varepsilon] \rightarrow 0$ as $N \rightarrow \infty$ by $\alpha > 1$, and $N^{d/2\alpha} f_N(x)^{-1} \rightarrow \infty$ as $N \rightarrow \infty$, and that $\mathbb{E}[|\bar{\omega}_x^\varepsilon|^\zeta] \leq C\varepsilon^{\zeta-\alpha} N^{\frac{d}{2\alpha}(\zeta-\alpha)}$ for the second term. Going back to (8.30), and using that $1 + z \leq e^z$ for any $z \in \mathbb{R}$, we get that the probability in the left-hand side of (8.30) is bounded by

$$\exp\left(-M_\varepsilon^{\frac{1}{2}(\alpha-1)} + C'M_\varepsilon^{2(\alpha-1)} \varepsilon^{\zeta-\alpha} N^{-\frac{d}{2}} \sum_{\|x\| \leq K\sqrt{N}} f_N(x)^\zeta\right) \leq c \exp\left(-M_\varepsilon^{\frac{1}{2}(\alpha-1)} + C''M_\varepsilon^{2(\alpha-1)} \varepsilon^{\zeta-\alpha}\right),$$

where we used again a sum integral comparison together with the fact that $\int_{\mathbb{R}^d} f(x)^a dx < +\infty$ for any $a < d/(d-2)$. Therefore, choosing $M_\varepsilon = \varepsilon^{-(\zeta-\alpha)/(\alpha-1)}$, the second term is a constant and we get that (8.30) is bounded by $\exp(-M_\varepsilon^{(\alpha-1)/2})$, which concludes the proof of (8.27). \square

8.4 Transversal fluctuations

The proof for transversal fluctuations are of order \sqrt{N} is a mimic of Subsection 7.3 and we omit it.

A Entropy-controlled Last-Passage Percolation (E-LPP)

Recall that Λ_r is the ball of radius r in \mathbb{R}^d , and that $\Lambda_r = \Lambda_r \cap \mathbb{Z}^d$. For any $m \in \mathbb{N}$, let $\Upsilon_m^{(r)}$ denote a set of m independent random variables $(Z_i)_{i=1}^m$ uniform in Λ_r , and let $\Upsilon_m^{(r)}$ denote a set of m distinct sites taken uniformly in Λ_r . We define the Entropy-controlled Last-Passage Percolation (E-LPP) as the maximal cardinality of a subset Δ of $\Upsilon_m^{(r)}$ or $\Upsilon_m^{(r)}$ with entropy smaller than B (recall the definition 4.1).

More precisely, we define the continuous and discrete E-LPP as follows

$$\begin{aligned} \mathcal{L}_m^{(B)}(r) &:= \max\{|\Delta| : \Delta \subset \Upsilon_m^{(r)}, \text{Ent}(\Delta) \leq B\}, \\ L_m^{(B)}(r) &:= \max\{|\Delta| : \Delta \subset \Upsilon_m^{(r)}, \text{Ent}(\Delta) \leq B\}. \end{aligned} \quad (A.1)$$

This continuous non-directed E-LPP has been considered in [9, Sec. 3], in dimension $d = 2$. The generalization of [9, Thm. 3.1] to the discrete E-LPP and to higher dimensions is straightforward (we give a brief proof below for the sake of completeness).

Theorem A.1. *There exists a constant $c_d > 0$ such that, for any $B > 0$ we have that for any $k \geq 1$*

$$\mathbb{P}\left(\mathcal{L}_m^{(B)}(r) > k\right) \leq \left(\frac{c_d B^{1/2} m^{1/d}}{rk}\right)^{dk}, \quad \mathbb{P}\left(L_m^{(B)}(r) > k\right) \leq \left(\frac{c_d B^{1/2} m^{1/d}}{rk}\right)^{dk}. \quad (\text{A.2})$$

We have the following corollary, whose proof is also given below.

Corollary A.2. *For any $b \leq d$, there exists a constant $c_{b,d}$, only depending on b and the dimension d , such that for any $B > 0$, any $m \geq 1$ and any r ,*

$$\mathbb{E}\left[\left(\frac{\mathcal{L}_m^{(B)}(r)}{(B^{1/2} m^{1/d}/r) \wedge m}\right)^b\right] \leq c_{b,d}, \quad \mathbb{E}\left[\left(\frac{L_m^{(B)}(r)}{(B^{1/2} m^{1/d}/r) \wedge m}\right)^b\right] \leq c_{b,d}. \quad (\text{A.3})$$

Proof of Theorem A.1. The proof relies on a first moment estimate. We prove the bound only in the continuous setting, the computations being the same in the discrete setting, using sum integral comparisons.

For any $B > 0$ we let $E_k(B) = \{\Delta = (x_i)_{i=1}^k : \text{Ent}(\Delta) \leq B\}$ and we denote \mathcal{N}_k the number of sets $\Delta \subset \Upsilon_m^{(r)}$ with $|\Delta| = k$, that have entropy at most B . Markov's inequality gives that

$$\mathbb{P}(\mathcal{L}_m^{(B)}(r) \geq k) = \mathbb{P}(\mathcal{N}_k \geq 1) \leq \mathbb{E}[\mathcal{N}_k].$$

By exchangeability of the points in $\Upsilon_m^{(r)}$, we get that

$$\mathbb{E}[\mathcal{N}_k] = \binom{m}{k} \mathbb{P}(\exists \sigma \in \mathfrak{S}_k \text{ s.t. } (Z_{\sigma(1)}, \dots, Z_{\sigma(k)}) \in E_k(B)) = \frac{m!}{(m-k)!} \frac{\text{Vol}(E_k(B))}{(c_d r^d)^k}, \quad (\text{A.4})$$

where we used that Z_1, \dots, Z_k are independent uniform random variables in the domain Λ_r , which has volume $c_d r^d$. Then, we can compute the volume of $E_k(B)$ by iteration, and we find that

$$\text{Vol}(E_k(B)) = \left(\frac{\pi^{d/2}}{\Gamma(\frac{d}{2} + 1)}\right)^k \frac{\Gamma(d)^k}{\Gamma(dk + 1)} B^{dk/2} \leq \left(\frac{C_d B^{d/2}}{k^d}\right)^k.$$

Going back to (A.4), and using that $\frac{m!}{(m-k)!} \leq m^k$, this concludes the proof. \square

Proof of Corollary A.2. We only prove the continuous case. The discrete case follows by the same argument. Note that if $m \leq B^{1/2} m^{1/d}/r$, the result is trivial. Hence, we assume that $m > B^{1/2} m^{1/d}/r$. Choose c_b such that $c_b \geq c_d k^{(b-1)/k}$ for all k , with c_d the constant in (A.2).

* If $c_b B^{d/2} m/r^d \leq 1/2$, then

$$\begin{aligned} \mathbb{E}\left[\left(\frac{\mathcal{L}_m^{(B)}(r)}{B^{1/2} m^{1/d}/r}\right)^b\right] &\leq \sum_{k=1}^{\infty} C k^{b-1} \mathbb{P}\left(\mathcal{L}_m^{(B)}(r) \geq k\right) \leq \sum_{k=1}^{\infty} C k^{b-1} \left(\frac{c_d B^{d/2} m}{r^d k^d}\right)^k \\ &\leq C \sum_{k=1}^{\infty} \left(c_b \frac{B^{d/2} m}{r^d k^d}\right)^k \leq 2C c_b \frac{B^{d/2} m}{r^d} \leq 2C \left(c_b \frac{B^{d/2} m}{r^d}\right)^{b/d}. \end{aligned} \quad (\text{A.5})$$

* If $c_b B^{d/2} m/r^d \geq 1/2$, note that $c_b u^{d/b} > c \vee 1$ for large u , then

$$\mathbb{E}\left[\left(\frac{\mathcal{L}_m^{(B)}(r)}{B^{1/2} m^{1/d}/r}\right)^b\right] = \int_0^{\infty} \mathbb{P}\left(\frac{\mathcal{L}_m^{(B)}(r)}{B^{1/2} m^{1/d}/r} \geq u^{1/b}\right) du \leq C + \int_C^{\infty} \left(c_d u^{-1/b}\right)^{du^{1/b} B^{1/2} m^{1/d}/r} du. \quad (\text{A.6})$$

Since $du^{1/b} B^{1/2} m^{1/d}/r \geq c'_d u^{1/b}$, we get that the last integral is bounded by a constant. \square

A.1 Applications of E-LPP

We now apply Theorem A.1 to prove Lemma 3.2 and Proposition 4.1.

Proof of Lemma 3.2.

Let us start by recalling Remark 3.1: if $s \in \mathcal{D}$ is such that $\text{Ent}(s) \leq 2$, then we have $s \in \mathcal{D}_2$. For any $m \in \mathbb{N}$, we denote $\Upsilon_m^{(2)} := \{\mathbf{Y}_i^{(2)}, 1 \leq i \leq m\}$, where we recall that $(\mathbf{Y}_i^{(2)})_{i \geq 1}$ are the positions of the order statistics of $(w, x) \in \mathcal{P}_2 = \{(w, x) \in \mathcal{P}, \|x\| \leq 2\}$ and are i.i.d. uniform random variables in Λ_2 . Recalling that the ordered statistics of weights $(\mathbf{M}_i^{(2)})_{i \geq 1}$ is a.s. decreasing, using (3.2) we have that

$$\pi_2(s) = \sum_{(w,x) \in \mathcal{P}_2} w \mathbb{1}_{\{x \in s([0,1])\}} = \sum_{j=0}^{\infty} \sum_{i=2^j}^{2^{j+1}} \mathbf{M}_i^{(2)} \mathbb{1}_{\{\mathbf{Y}_i^{(2)} \in s([0,1])\}} \leq \sum_{j=0}^{\infty} \mathbf{M}_{2^j}^{(2)} \mathcal{L}_{2^{j+1}}^{(2)}(2), \quad (\text{A.7})$$

where $\mathcal{L}_m^{(2)}(2)$ is the continuous E-LPP defined above in (A.1). Let $\delta > 0$ be such that $\frac{1}{\alpha} - \frac{1}{d} > \delta$. A union bound gives that for any $t > 0$,

$$\mathbb{P}\left(\sup_{s: \text{Ent}(s) < 2} \{\pi_2(s)\} > t\right) \leq \sum_{j=0}^{\infty} \mathbb{P}\left(\mathbf{M}_{2^j}^{(2)} \mathcal{L}_{2^{j+1}}^{(2)}(2) > Ct 2^{j(\frac{1}{d} - \frac{1}{\alpha} + \delta)}\right). \quad (\text{A.8})$$

where $1/C = \sum_{j \in \mathbb{N}} 2^{j(\frac{1}{d} - \frac{1}{\alpha} + \delta)}$. Then, (A.8) is smaller than

$$\sum_{j=0}^{\infty} \left[\mathbb{P}\left(\mathcal{L}_{2^{j+1}}^{(2)}(2) > C_0 \log_2(2+t) 2^{(j+1)/d}\right) + \mathbb{P}\left(\mathbf{M}_{2^j}^{(2)} > C_1 \frac{t}{\log_2(2+t)} 2^{j(-\frac{1}{\alpha} + \delta)}\right) \right], \quad (\text{A.9})$$

where C_0 is a constant chosen (large enough) below, and C_1 depends on C and C_0 . We use (A.2) with $m = 2^{j+1}$ and $k = C_0 \log_2 t 2^{(j+1)/d}$. Then if C_0 is taken large enough, we have

$$\mathbb{P}\left(\mathcal{L}_{2^{j+1}}^{(2)}(2) > C_0 \log_2(2+t) 2^{(j+1)/d}\right) \leq \left(\frac{1}{2}\right)^{dC_0 \log_2(2+t) 2^{(j+1)/d}}.$$

For the second term in (A.9), we recall that $\mathbf{M}_i^{(2)} \stackrel{(d)}{=} (2\sqrt{2})^{d/\alpha} \text{Gamma}(i)^{-1/\alpha}$, so that $\mathbb{E}[(\mathbf{M}_i^{(2)})^\alpha] \leq ci^{-1}$, where c is a constant independent of i . Therefore, using Markov's inequality we get that

$$\mathbb{P}\left(\mathbf{M}_{2^j}^{(2)} > C_1 \frac{t}{\log_2(2+t)} 2^{j(-\frac{1}{\alpha} + \delta)}\right) \leq c \log_2(2+t)^\alpha t^{-\alpha} 2^{-j\delta\alpha}.$$

All together, we get that

$$\begin{aligned} \mathbb{P}\left(\sup_{s: \text{Ent}(s) < 2} \{\pi_2(s)\} > t\right) &\leq \sum_{j=0}^{\infty} \left(c \log_2(2+t)^\alpha t^{-\alpha} 2^{-j\delta\alpha} + 2^{-dC_0(\log_2(2+t))2^{(j+1)/d}} \right) \\ &\leq c' \log_2(2+t)^\alpha t^{-\alpha} + c' t^{-c''C_0} \leq c''' \log_2(2+t)^\alpha t^{-\alpha}, \end{aligned}$$

where the constant c'' depends only on d , and C_0 has been fixed large enough so that $c''C_0 > \alpha$. This concludes the proof, by replacing the exponent α by some $a < \alpha$ and getting rid of the term $\log_2(2+t)^\alpha$ by adjusting the constant. \square

Proof of Proposition 4.1.

To prove Proposition 4.1, we start by bounding the tail of $\Omega_r^{(\ell)}(\Delta)$, $\Omega_r^{(>\ell)}(\Delta)$ in Lemmas A.3 and A.4 respectively.

Lemma A.3. Let $\alpha \in (0, d)$. There exists a constant c , such that for any $B > 0$, $r > 0$ and $t > 1$,

$$\mathbb{P} \left(\sup_{\text{Ent}(\Delta) \leq B} \Omega_r^{(\ell)}(\Delta) \geq t \times r^{\frac{d}{\alpha}-1} B^{1/2} \right) \leq ct^{-\alpha d/(\alpha+d)}.$$

Proof. The proof is the discrete version of Lemma 3.2 proven above: using the discrete E-LPP defined in (A.1), we can write

$$\sup_{\text{Ent}(\Delta) \leq B} \Omega_r^{(\ell)}(\Delta) \leq \sum_{j=0}^{\log_2 \ell} \sum_{i=2^j}^{2^{j+1}-1} M_i^{(r)} \mathbb{1}_{\{Y_i^{(r)} \in \Delta\}} \leq \sum_{j=0}^{\log_2 \ell} M_{2^j}^{(r)} L_{2^{j+1}}^{(B)}(r). \quad (\text{A.10})$$

The proof then follows the same lines as above, but we keep some details to help keep track of the different parameters. Let $\delta > 0$ be such that $1/\alpha - 1/d > 2\delta$ and let $1/C = \sum_{j=0}^{\infty} 2^{j(1/d-1/\alpha+2\delta)}$. For $\varepsilon \in [0, 1]$, by two consecutive union bounds we get

$$\begin{aligned} \mathbb{P} \left(\sup_{\text{Ent}(\Delta) \leq B} \Omega_r^{(\ell)}(\Delta) \geq t \times r^{\frac{d}{\alpha}-1} B^{1/2} \right) &\leq \sum_{j=0}^{\log_2 \ell} \mathbb{P} \left(\frac{M_{2^j}^{(r)}}{r^{d/\alpha}} \frac{L_{2^{j+1}}^{(B)}(r)}{B^{1/2}/r} \geq Ct 2^{j(\frac{1}{d}-\frac{1}{\alpha}+2\delta)} \right) \\ &\leq \sum_{j=0}^{\log_2 \ell} \mathbb{P} \left(\frac{M_{2^j}^{(r)}}{r^{d/\alpha}} > C_1 t^{1-\varepsilon} 2^{-\frac{j}{\alpha}+j\delta} \right) + \sum_{j=0}^{\log_2 \ell} \mathbb{P} \left(\frac{L_{2^{j+1}}^{(B)}(r)}{B^{\frac{1}{2}}/r} > C_2 t^\varepsilon (2^{j+1})^{\frac{1}{d}+\delta} \right). \end{aligned} \quad (\text{A.11})$$

For the first term above, we can use Lemma 5.1 in [8] to get that for any $u > 0$ and $\ell \leq |\Lambda_r|$, we have

$$\mathbb{P} \left(M_\ell^{(r)} > u r^{d/\alpha} \ell^{-1/\alpha} \right) < (cu)^{-\alpha\ell}. \quad (\text{A.12})$$

For the second term, we apply Markov's inequality together with Corollary A.2 with $b = d$. All together, we get that the left-hand side of (A.11) is bounded by

$$\sum_{j=0}^{\infty} (ct^{1-\varepsilon})^{-\alpha 2^j} 2^{-j 2^j \delta \alpha} + \sum_{j=0}^{\infty} c' (C_2 2^{(j+1)\delta} t^\varepsilon)^{-d} \leq C' t^{(1-\varepsilon)\alpha} + C' t^{-\varepsilon d}. \quad (\text{A.13})$$

Finally, we choose $\varepsilon = \alpha/(\alpha + d)$ and the proof is completed. \square

Lemma A.4. For any $\ell \leq |\Lambda_r|$, there exists a constant c , such that for any $B > 0$, $r > 0$ and $t > 1$,

$$\mathbb{P} \left(\sup_{\text{Ent}(\Delta) \leq B} \Omega_r^{(>\ell)} \geq t \times r^{\frac{d}{\alpha}-1} \ell^{\frac{1}{d}-\frac{1}{\alpha}} B^{1/2} \right) \leq ct^{-\frac{\alpha\ell d}{\alpha\ell+d}}.$$

Proof. The proof is similar to that of the previous lemma. Similarly to (A.10), we can write

$$\sup_{\text{Ent}(\Delta) \leq B} \Omega_r^{(>\ell)}(\Delta) \leq \sum_{j=0}^{\log_2(|\Lambda_r|/\ell)} M_{2^j \ell}^{(r)} L_{2^{j+1} \ell}^{(B)}(r).$$

Using (A.12), we have

$$\mathbb{P} \left(\frac{M_{2^j \ell}^{(r)}}{r^{d/\alpha} \ell^{1/\alpha}} > C_1 t^{1-\varepsilon} (2^j)^{-\frac{1}{\alpha}+\delta} \right) \leq (ct^{1-\varepsilon})^{-2^j \alpha \ell} 2^{-j 2^j \alpha \delta \ell} \quad (\text{A.14})$$

and by Markov's inequality and Corollary A.2

$$\mathbb{P} \left(\frac{L_{2^{j+1} \ell}^{(B)}(r)}{\ell^{\frac{1}{d}} B^{\frac{1}{2}}/r} > C_2 t^\varepsilon (2^{j+1})^{\frac{1}{d}+\delta} \right) \leq (C_2 2^{(j+1)\delta})^{-d} t^{-\varepsilon d}. \quad (\text{A.15})$$

The result follows by a similar bound to (A.13). \square

Proof of Proposition 4.1. Let us introduce

$$\mathcal{T}_{N,r}^{\beta,(\ell)}([a,b)) := \max_{\Delta \subset \Lambda_r, \text{Ent}(\Delta) \in [a,b)} \left\{ \beta \Omega_r^{(\ell)}(\Delta) - \text{Ent}_N(\Delta) \right\}. \quad (\text{A.16})$$

For any $b > 0$, we can then decompose the variational problem as

$$\begin{aligned} \mathcal{T}_{N,r}^{\beta,(\ell)} &= \max \left\{ \mathcal{T}_{N,r}^{\beta,(\ell)}([0,b)), \sup_{k \geq 0} \left\{ \mathcal{T}_{N,r}^{\beta,(\ell)}([2^k b, 2^{k+1} b)) \right\} \right\} \\ &\leq \max_{\Delta \subset \Lambda_r, \text{Ent}(\Delta) \leq b} \beta \Omega_r^{(\ell)}(\Delta) \vee \sup_{k \geq 0} \left\{ \beta \max_{\Delta \subset \Lambda_r, \text{Ent}(\Delta) \leq 2^{k+1} b} \Omega_r^{(\ell)} - \frac{2^k b}{N} \right\} \end{aligned}$$

Choosing $b = t(N\beta r^{\frac{d}{\alpha}-1})^2$ and applying a union bound, we get

$$\mathbb{P} \left(\mathcal{T}_{N,r}^{\beta,(\ell)} \geq tN \times (\beta r^{\frac{d}{\alpha}-1})^2 \right) \leq \sum_{k=0}^{\infty} \mathbb{P} \left(\max_{\Delta \subset \Lambda_r, \text{Ent}(\Delta) \leq 2^k t(N\beta r^{\frac{d}{\alpha}-1})^2} \Omega_r^{(\ell)} \geq 2^{k-1} tN \beta (r^{\frac{d}{\alpha}-1})^2 \right)$$

Therefore, applying Lemma A.3 (with $B = 2^k t(N\beta r^{\frac{d}{\alpha}-1})^2$ and $t' = 2^{k/2-1} t^{1/2}$), we obtain that

$$\mathbb{P} \left(\mathcal{T}_{N,r}^{\beta,(\ell)} \geq tN \times (\beta r^{\frac{d}{\alpha}-1})^2 \right) \leq c \sum_{k=0}^{+\infty} t^{-\frac{\alpha d}{2(\alpha+d)}} 2^{-\frac{\alpha d}{2(\alpha+d)} k} \leq c' t^{-\frac{\alpha d}{2(\alpha+d)}}. \quad (\text{A.17})$$

For (4.9) the proof follows the same lines as (A.16)-(A.17), by applying Lemma A.4 instead of Lemma A.3. \square

Let us now prove another technical result: we show that the paths that maximize $\widehat{\mathcal{T}}_\beta$ are concentrated around the argmax of the variational problem, which is needed to get (8.1) — this is the analogous of Lemma 4.1 in [2].

Lemma A.5. *On the event $\{\widehat{\mathcal{T}}_\beta = 0\}$, for any $\eta \in (0, 1)$ we have that*

$$\widehat{\mathcal{T}}_{\beta, \geq \eta} := \sup_{s \in \mathcal{D}_1, \sup \|s\| \geq \eta, \widehat{\text{Ent}}(s) < +\infty} \left\{ \pi(s) - \frac{1}{\beta} \widehat{\text{Ent}}(s) \right\} < 0, \quad a.s. \quad (\text{A.18})$$

Proof. Recall the definition (3.1) of \mathcal{D}_1 and (3.4) of $\pi_1^{(\ell)}$ (recall that $\widehat{\text{Ent}}(s) = +\infty$ if $\sup_{[0,1]} \|s\| > 1$). For $\ell \geq 1$, we define

$$\rho^{(\ell)} := \sup_{s \in \mathcal{D}_1} |\pi(s) - \pi^{(\ell)}(s)|, \quad \widehat{\mathcal{T}}_{\beta, \geq \eta}^{(\ell)} := \sup_{s \in \mathcal{D}_1, \sup \|s\| \geq \eta, \widehat{\text{Ent}}(s) < +\infty} \left\{ \pi_1^{(\ell)}(s) - \frac{1}{\beta} \widehat{\text{Ent}}(s) \right\}.$$

For all ℓ we have that $\widehat{\mathcal{T}}_{\beta, \geq \eta} \leq \widehat{\mathcal{T}}_{\beta, \geq \eta}^{(\ell)} + \rho^{(\ell)}$, so the proof is a consequence of the following:

$$\lim_{\ell \rightarrow +\infty} \rho^{(\ell)} = 0, \quad a.s., \quad \limsup_{\ell \rightarrow +\infty} \widehat{\mathcal{T}}_{\beta, \geq \eta}^{(\ell)} < 0 \quad a.s. \quad (\text{A.19})$$

To prove the first part of (A.19), reasoning as in the proof of (4.13) in [8], an integration by part gives that

$$\pi(s) - \pi^{(\ell)}(s) \leq \sum_{i=\ell+1}^{\infty} \mathcal{L}_i(\mathbf{M}_i - \mathbf{M}_{i+1}) + \limsup_{n \rightarrow \infty} \mathbf{M}_n \mathcal{L}_n. \quad (\text{A.20})$$

where $\mathbf{M}_i := \mathbf{M}_i^{(1)}$ and $\mathcal{L}_n := \mathcal{L}_n^{(d/2)}(1)$ are defined in (3.3) and (A.1) respectively; note that the restriction $\widehat{\text{Ent}}(s) < +\infty$ implies that s has length at most 1, so $\text{Ent}(s) \leq d/2$. The law of large numbers gives $\lim_{n \rightarrow \infty} n^{1/\alpha} \mathbf{M}_n = c_1$ a.s., and Theorem A.1 gives

$\limsup_{n \rightarrow \infty} n^{-1/d} \mathcal{L}_n < +\infty$ a.s. Since $\alpha < d$, we therefore conclude that $\limsup_{n \rightarrow \infty} \mathbf{M}_n \mathcal{L}_n = 0$ a.s.

To complete the proof of the first part of (A.19), we let $U_\ell := \sum_{i>\ell} \mathcal{L}_i(\mathbf{M}_i - \mathbf{M}_{i+1})$ and we show that $U_\ell < +\infty$ a.s., by showing that $\mathbb{E}[U_\ell^2]$ is finite. For any $\varepsilon > 0$, by Cauchy-Schwarz inequality we have that

$$U_\ell \leq \left(\sum_{i>\ell} (i^{-\frac{1}{2}-\varepsilon})^2 \right)^{1/2} \left(\sum_{i>\ell} (i^{\frac{1}{2}+\varepsilon} \mathcal{L}_i(\mathbf{M}_i - \mathbf{M}_{i+1}))^2 \right)^{1/2}.$$

Then, we get that for ℓ large enough

$$\mathbb{E}[U_\ell^2] \leq C \sum_{i>\ell} i^{1+2\varepsilon} \mathbb{E}[(\mathcal{L}_i)^2] \mathbb{E}[(\mathbf{M}_i - \mathbf{M}_{i+1})^2] \leq C' \sum_{i>\ell} i^{2\varepsilon+2/d-2/\alpha-1} < +\infty,$$

where we used Corollary A.2 and the fact that $\mathbb{E}[(\mathbf{M}_i - \mathbf{M}_{i+1})^2] \leq ci^{-2-2/\alpha}$ (see for instance Equation (7.2) in [24]). Provided ε is small enough so that $\varepsilon + 1/d - 1/\alpha < 0$, we obtain that $\mathbb{E}[U_\ell^2] < \infty$ so $U_\ell < \infty$ a.s. We therefore get that $\lim_{\ell \rightarrow \infty} U_\ell = 0$, which proves the first part of (A.19).

We prove the second part of (A.19) by contradiction. Note that by (1.9)-(1.10), if $\widehat{\text{Ent}}(s) < +\infty$, there exists some parametrization $\varphi \in \Phi$ such that $(s \circ \varphi)(t)$ is 1-Lipschitz on $[0, 1]$ and $\pi(s) = \pi(s \circ \varphi)$ for all $\varphi \in \Phi$. Hence, it is enough to consider only 1-Lipschitz paths $\{s \in \mathcal{D}_1 : \|s'\| \leq 1\}$. Let us suppose that there exists a sequence $\tilde{s}^{(\ell)}$ such that $\sup \|\tilde{s}^{(\ell)}\| \geq \eta$ and $\limsup \mathcal{Z}_\beta^{(\ell)}(\tilde{s}^{(\ell)}) \geq 0$, where we set $\mathcal{Z}_\beta^{(\ell)}(s) := \pi^{(\ell)}(s) - \frac{1}{\beta} \widehat{\text{Ent}}(s)$ (and $\mathcal{Z}_\beta(s)$ when $\ell = +\infty$).

We observe that $s \mapsto \mathcal{Z}_\beta^{(\ell)}(s)$ is upper semi-continuous because $\widehat{\text{Ent}}(s)$ is lower semi-continuous (it is a rate function of a large deviation principle) and $s \mapsto \pi^{(\ell)}(s)$ is upper semi-continuous by construction (we refer to Lemma 4.4 in [8] for more details). Since the set of 1-Lipshitz functions on $[0, 1]$ is compact for the uniform norm by Ascoli-Arzelà theorem, we can suppose that $\tilde{s}^{(\ell)}$ converges uniformly to \tilde{s} as $\ell \rightarrow +\infty$. Then, using the upper semi-continuity we get

$$\widehat{\mathcal{T}}_\beta \geq \mathcal{Z}_\beta(\tilde{s}) \geq \limsup_{\ell \rightarrow +\infty} \mathcal{Z}_\beta(\tilde{s}^{(\ell)}) \geq \limsup_{\ell \rightarrow +\infty} \mathcal{Z}_\beta^{(\ell)}(\tilde{s}^{(\ell)}) \geq 0.$$

Using that when $\widehat{\mathcal{T}}_\beta = 0$ the maximiser is unique and is given by $\widehat{s}_\beta \equiv 0$, we conclude that $\tilde{s} = 0$, which contradicts the fact that $\sup \|\tilde{s}\| \geq \eta$. \square

Remark A.6. In general, we can define

$$\widehat{\mathcal{T}}_{\beta, \|s - \widehat{s}_\beta\| \geq \eta} := \sup_{s \in \mathcal{D}, \sup \|s - \widehat{s}_\beta\| \geq \eta, \widehat{\text{Ent}}(s) < +\infty} \left\{ \pi(s) - \frac{1}{\beta} \widehat{\text{Ent}}(s) \right\},$$

where \widehat{s}_β is the maximiser of $\widehat{\mathcal{T}}_\beta$; and similarly for $\mathcal{T}_{\beta, \|s - s_\beta\| \geq \eta}$ with s_β the maximizer of \mathcal{T}_β . This requires the existence and uniqueness of the maximisers $\widehat{s}_\beta, s_\beta$, which follows similarly to Section 4.6 in [8]. Then, one has the analogous of Lemma A.5, that is $\widehat{\mathcal{T}}_{\beta, \|s - \widehat{s}_\beta\| \geq \eta} < \widehat{\mathcal{T}}_\beta$ and $\mathcal{T}_{\beta, \|s - s_\beta\| \geq \eta} < \mathcal{T}_\beta$ a.s.

B Simple random walk estimates

Let us collect some technical results on d -dimensional simple random walks that we use in the article.

B.1 Large and moderate deviations for the simple random walk

Lemma B.1. Let $(S_n)_{n \geq 0}$ be a simple symmetric random walk on \mathbb{Z}^d . Let $\xi \in (\frac{1}{2}, 1]$ be fixed. Then, for any $x \in \mathbb{R}^d$, denote $x_N^{(\xi)}$ the element of \mathbb{Z}^d with the same parity as N closest to xN^ξ . Then we have the local large deviation result.

$$\lim_{N \rightarrow \infty} -\frac{1}{N^{2\xi-1}} \log \mathbf{P}(S_N = x_N^{(\xi)}) = \begin{cases} J_d(x) & \text{if } \xi = 1, \\ \frac{d}{2} \|x\|^2 & \text{if } \xi \in (\frac{1}{2}, 1), \end{cases}$$

where $J_d(\cdot)$ is a given rate function, which verifies $J(\|x\|_1) \leq J_d(x) \leq J(\|x\|_1) + \log d$, where $J(t) = \frac{1}{2}(1+t) \log(1+t) + \frac{1}{2}(1-t) \log(1-t)$ is the large deviation rate function for the random walk in dimension $d = 1$; in particular, we have $J_d(x) < +\infty$ if $\|x\|_1 \leq 1$ and $J_d(x) = +\infty$ if $\|x\|_1 > 1$. We also have $J_d(x) \geq \frac{1}{2} \|x\|^2$ for all $x \in \mathbb{R}^d$, and $J_d(x) \sim \frac{d}{2} \|x\|^2$ as $\|x\| \downarrow 0$.

Proof. This is a standard result, that directly derives from local large deviation results for the simple random walk in dimension $d = 1$, by decomposing over the number of steps in each direction. Let us just treat the case of $\xi = 1$, the case $\xi \in (\frac{1}{2}, 1)$ being analogous. The rate function J_d can be related to the rate function J in the following way, decomposing over the proportion of the time spent in each direction

$$J_d(x) = \inf_{\substack{u_1, \dots, u_d \in [0,1] \\ u_1 + \dots + u_d = 1}} \sum_{i=1}^d u_i \left(J\left(\frac{x_i}{u_i}\right) + \log(du_i) \right), \quad (\text{B.1})$$

where the term $\sum_{i=1}^d u_i \log(du_i)$ comes from the entropic cost of spending a proportion u_i of the time in direction i ; more precisely, we have $\lim_{N \rightarrow \infty} -\frac{1}{N} \log \left(\frac{N!}{(u_1 N)! \dots (u_d N)!} d^{-N} \right) = \sum_{i=1}^d u_i \log(du_i)$.

The rate function J_d does not appear to have a nice expression, but some properties can be derived. Note that $J(\cdot)$ being an even function we can replace $J(x_i/u_i)$ by $J(|x_i|/u_i)$ in the above expression. Since $J(\cdot)$ is a convex function we get that

$$\sum_{i=1}^d u_i J\left(\frac{|x_i|}{u_i}\right) \geq J\left(\sum_{i=1}^d |x_i|\right) = J(\|x\|_1).$$

Hence, using also that $\sum_{i=1}^d u_i \log(du_i) \geq 0$ for all $u_1, \dots, u_d \in [0, 1]$ with $u_1 + \dots + u_d = 1$, we get that $J_d(x) \geq J(\|x\|_1)$. On the other hand, since $\sum_{i=1}^d u_i \log(du_i) \leq \log d$, taking $u_i = |x_i|/\|x\|_1$ we obtain that $J_d(x) \leq J(\|x\|_1) + \log d$.

To prove the last inequality, let us simply notice that $J(t) \geq \frac{1}{2}t^2$ for all $t \in \mathbb{R}$, so we get that $J_d(x) \geq \frac{1}{2} \|x\|_1^2 \geq \frac{1}{2} \|x\|^2$. Also, notice that when $\|x\| \downarrow 0$, we have that $J_d(x) \rightarrow 0$: for instance, taking $u_i = \frac{1}{d}$ in (B.1), we get that $J_d(x) \leq \frac{1}{d} \sum_{i=1}^d J(dx_i)$. Since the infimum in (B.1) goes to 0, it means that it is attained for some u_i approaching $\frac{1}{d}$ (so that $\log(du_i)$ goes to 0), and we therefore get that $J_d(x) \sim \frac{1}{d} \sum_{i=1}^d J(dx_i)$ as $\|x\| \downarrow 0$. Using that $J(t) \sim \frac{1}{2}t^2$ as $t \downarrow 0$, we get that $J_d(x) \sim \frac{d}{2} \|x\|^2$. \square

B.2 Probabilities that a given set of points is visited

Here, for clarity, we separate the cases $\xi \in (\frac{1}{2}, 1)$ and $\xi = 1$. For a set of (ordered) points $\Delta = (y_1, \dots, y_k) \in \mathbb{Z}^d$, with some abuse of notation we write $\Delta \subset \mathcal{R}_N$ to mean that the points of Δ are visited in their order by the random walk before time N : put otherwise

$$\{\Delta \subset \mathcal{R}_N\} = \{\exists 0 < t_1 < \dots < t_k \leq N, \text{ s.t. } S_{t_1} = y_1, \dots, S_{t_k} = y_k\}. \quad (\text{B.2})$$

We now state large deviation principles for events $\{\Delta \subset \mathcal{R}_N\}$.

Lemma B.2. Let $\xi \in (\frac{1}{2}, 1)$. For any ordered set $\Delta = (x_1, \dots, x_k) \subset \mathbb{R}^d$ of distinct points, we have (omitting integer parts for notational simplicity)

$$\lim_{N \rightarrow \infty} -\frac{1}{N^{2\xi-1}} \log \mathbf{P}((x_1 N^\xi, \dots, x_k N^\xi) \subset \mathcal{R}_N) = \text{Ent}(\Delta),$$

where the entropy $\text{Ent}(\Delta)$ is defined in (4.1).

Proof. We only treat the case $k = 2$ to lighten notation; the general case is analogous.

Lower bound: By a local large deviation, see e.g. [30, Theorem 3], we get

$$\begin{aligned} \mathbf{P}(\exists 1 \leq s < t \leq N, \text{ s.t. } S_s = xN^\xi, S_t = yN^\xi) &\geq \max_{1 \leq s < N} \mathbf{P}(S_s = xN^\xi, S_N = yN^\xi) \\ &= \max_{0 < u \leq 1} \frac{c_d}{(u(1-u))^{\frac{d}{2}} N^d} \exp\left(-N^{2\xi-1} \frac{d}{2} \left(\frac{\|x\|^2}{u} + \frac{\|y-x\|^2}{1-u} + o(1)\right)\right) \end{aligned}$$

Now the term $\frac{\|x\|^2}{u} + \frac{\|y-x\|^2}{1-u}$ is minimal for $u = \frac{\|x\|}{\|x\| + \|y-x\|}$ and choosing that specific u gives that

$$\mathbf{P}(\exists 1 \leq s < t \leq N, \text{ s.t. } S_s = xN^\xi, S_t = yN^\xi) \geq \frac{c_{d,x,y}}{N^d} e^{-N^{2\xi-1}(1+o(1))\text{Ent}(\Delta)},$$

which proves the lower bound. Note that if $\|x\|$ and $\|y-x\|$ are bounded away from 0 and ∞ , then so is $c_{d,x,y}$.

Upper bound: Using again [30, Theorem 3], we have that

$$\begin{aligned} \mathbf{P}(\exists 1 \leq s < t \leq N, \text{ s.t. } S_s = xN^\xi, S_t = yN^\xi) &\leq \sum_{s=1}^N \sum_{t=s+1}^N \mathbf{P}(S_s = xN^\xi, S_t = yN^\xi) \\ &\leq \sum_{s=1}^N \sum_{t=s+1}^N \frac{c_d}{(s(t-s))^{\frac{d}{2}}} \exp\left(-N^{2\xi-1} \frac{d}{2} \left(\frac{\|x\|^2}{s/N} + \frac{\|y-x\|^2}{(t-s)/N} + o(1)\right)\right), \end{aligned} \quad (\text{B.3})$$

where the $o(1)$ is uniform in N . To get an upper bound for (B.3), we minimize the quantity in the exponent, which is $N^{2\xi-1}\text{Ent}(\Delta)$ (as above). Then for any $d \geq 1$, (B.3) is bounded above by a constant times

$$e^{-(1+o(1))N^{2\xi-1}\text{Ent}(\Delta)} \sum_{s=1}^N \sum_{t=s+1}^N \frac{1}{(s(t-s))^{d/2}} \leq cN e^{-(1+o(1))N^{2\xi-1}\text{Ent}(\Delta)},$$

which gives the upper bound. \square

Let us state the analogous result in the case $\xi = 1$. We do not prove it, since it is analogous to Lemma B.2.

Lemma B.3. For any ordered set $\Delta = (x_1, \dots, x_k) \subset \mathbb{R}^d$ of distinct points, we have (omitting integer parts for notational simplicity)

$$\lim_{N \rightarrow \infty} -\frac{1}{N} \log \mathbf{P}((x_1 N, \dots, x_k N) \subset \mathcal{R}_N) = \widehat{\text{Ent}}(\Delta),$$

where the entropy $\widehat{\text{Ent}}(\Delta)$ is defined in (4.10).

Notice here that the definition (4.10) of the entropy $\widehat{\text{Ent}}(\Delta)$ includes the infimum over the choices of times $0 < t_1 < \dots < t_k \leq 1$ at which the random walk visits the points.

To conclude, we give a uniform upper bound on the probability that a given set is visited.

Lemma B.4. *There are positive constants $C_1 := C_1(d)$ and $C_2 := C_2(d)$ such that for any set $\Delta = (x_1, \dots, x_\ell) \subset \mathbb{Z}^d$ of distinct points, we have in dimension $d \geq 2$*

$$\mathbf{P}(\Delta \subset \mathcal{R}_N) \leq (C_1)^\ell e^{-\frac{C_2}{N} \text{Ent}(\Delta)},$$

where $\text{Ent}_N(\Delta)$ is defined in (4.2).

Proof. We write $x_0 = 0$ and $k_0 = 0$ by convention. Let us start with the case of dimension $d \geq 3$. By a union bound (recalling (B.2)), we have

$$\mathbf{P}((x_1, \dots, x_\ell) \subset \mathcal{R}_N) \leq \sum_{k_1=1}^N \cdots \sum_{k_\ell=k_{\ell-1}+1}^N \prod_{i=1}^{\ell} \mathbf{P}(S_{k_i} - S_{k_{i-1}} = x_i - x_{i-1}).$$

Now, by standard local large deviation results, there exist constants c_1, c_2 such that for any $t \geq 1$ and $x \in \mathbb{Z}^d$,

$$\mathbf{P}(S_t = x) \leq \frac{c_1}{t^{d/2}} e^{-c_2 \|x\|^2/t}.$$

Thus, since for all $0 = u_0 < u_1 < \dots < u_\ell \leq 1$ we have

$$\sum_{i=1}^{\ell} \frac{\|x_i - x_{i-1}\|^2}{u_i - u_{i-1}} \geq \text{Ent}(\Delta), \quad (\text{B.4})$$

we get that

$$\mathbf{P}((x_1, \dots, x_\ell) \subset \mathcal{R}_N) \leq c_1^\ell e^{-\frac{C_2}{N} \text{Ent}(\Delta)} \sum_{k_1=1}^N \cdots \sum_{k_\ell=k_{\ell-1}+1}^N \prod_{i=1}^{\ell} \frac{1}{(k_i - k_{i-1})^{d/2}} \leq (C_1)^\ell e^{-\frac{C_2}{N} \text{Ent}(\Delta)},$$

which concludes the proof in dimension $d \geq 3$.

For the case of the dimension $d = 2$, instead of the union bound, we use that, by Markov's property

$$\mathbf{P}((x_1, \dots, x_\ell) \subset \mathcal{R}_N) \leq \sum_{k_1=1}^N \cdots \sum_{k_\ell=k_{\ell-1}+1}^N \prod_{i=1}^{\ell} \mathbf{P}(H_{x_i - x_{i-1}} = k_i - k_{i-1}),$$

where $H_x = \min\{t \geq 1, S_t = x\}$ is the hitting time of the site x . Now, from Uchiyama's [32, Thm. 1.4], we get that uniformly for $x \in \mathbb{Z}^d \setminus \{0\}$

$$\mathbf{P}(H_x = k) \leq C \frac{\log(1 + \|x\|)}{k(\log(1 + k))^2} e^{-c\|x\|^2/k} \mathbb{1}_{\{k \geq c\|x\|\}}.$$

Using again (B.4), we get

$$\mathbf{P}((x_1, \dots, x_\ell) \subset \mathcal{R}_N) \leq c_1^\ell e^{-\frac{C_2}{N} \text{Ent}(\Delta)} \prod_{i=1}^{\ell} \left(\sum_{k_i=k_{i-1}+c\|x_i-x_{i-1}\|}^N \frac{\log(1 + \|x_i - x_{i-1}\|)}{(k_i - k_{i-1}) \log(1 + k_i - k_{i-1})^2} \right).$$

All the sums are bounded, so this concludes the proof in the case of the dimension $d = 2$. \square

Intersection of ranges of independent random walks

Recall that $J_N := \sum_{x \in \mathbb{Z}^d} \mathbf{P}(x \in \mathcal{R}_N)^2$.

Lemma B.5. *We have the following asymptotics, as $N \rightarrow +\infty$,*

$$J_N = (1 + o(1)) \begin{cases} c_2 \frac{N}{(\log N)^2} & \text{if } d = 2, \\ c_3 \sqrt{N} & \text{if } d = 3, \end{cases} \quad J_N = (1 + o(1)) \begin{cases} c_4 \log N & \text{if } d = 4, \\ c_d & \text{if } d = 5. \end{cases}$$

Proof. First of all, we can rewrite $J_N = \mathbf{E}^{\otimes 2}[\mathcal{J}_N]$, with $\mathcal{J}_N := |\mathcal{R}_N^{(1)} \cap \mathcal{R}_N^{(2)}|$ the intersection of ranges $\mathcal{R}_N^{(1)}$ and $\mathcal{R}_N^{(2)}$ of two independent random walks $S^{(1)}$ and $S^{(2)}$. In Chen's book [16], the weak convergence of \mathcal{J}_N is considered: in dimension $d = 2$, $\frac{(\log N)^2}{N} \mathcal{J}_N$ is shown to converge in distribution, see [16, Thm. 5.3.4]; in dimension $d = 3$, $\frac{1}{\sqrt{N}} \mathcal{J}_N$ is shown to converge in distribution, see [16, Thm. 5.3.4]; in dimension $d = 4$, $\frac{1}{\log N} \mathcal{J}_N$ is shown to converge in distribution, see [16, Thm. 5.5.1]. The convergence of the expectation $J_N = \mathbb{E}[\mathcal{J}_N]$ then comes as a consequence of the uniformity in the integrability of \mathcal{J}_N , see [16, § 6.2].

In dimension $d = 5$, \mathcal{J}_N converges to $|\mathcal{R}_\infty^{(1)} \cap \mathcal{R}_\infty^{(2)}|$, which is a.s. finite. The convergence of its expectation J_N follows by monotone convergence, the fact that $\mathbb{E}[|\mathcal{R}_\infty^{(1)} \cap \mathcal{R}_\infty^{(2)}|] < +\infty$ being given in (3.19) (with $\alpha = 2$). \square

C Technical estimate on the environment

We collect here some estimates that are needed along the paper.

Lemma C.1. *For any non-negative integer $p < \alpha$ and any sequence $k_N \geq 1$, we have:*

- If $\beta_N k_N \geq 1$,

$$\left| \mathbb{E}[\exp(\beta_N(\omega - \mu) \mathbb{1}_{\{\omega \leq k_N\}})] - 1 - \sum_{i=1}^p \frac{\beta_N^i}{i!} \mathbb{E}[(\omega - \mu)^i] \right| \leq C e^{\beta_N k_N} \begin{cases} \beta_N^\alpha & \text{if } p+1 > \alpha, \\ \beta_N^\alpha \log k_N & \text{if } p+1 = \alpha, \\ \beta_N^{p+1} & \text{if } p+1 < \alpha; \end{cases} \quad (\text{C.1})$$

- If $\beta_N k_N < 1$, in the case $\alpha < p+1$, we get

$$\left| \mathbb{E}[\exp(\beta_N(\omega - \mu) \mathbb{1}_{\{\omega \leq k_N\}})] - 1 - \sum_{i=1}^p \frac{\beta_N^i}{i!} \mathbb{E}[(\omega - \mu)^i] \right| \leq C \beta_N k_N^{1-\alpha} \quad (\text{C.2})$$

with an extra factor $\log k_N$ in the upper bound if $\alpha = p+1$.

Proof. We can use that for any $x \in \mathbb{R}$, we have $|e^x - \sum_{i=0}^p \frac{x^i}{i!}| \leq |x|^{p+1} e^{|x|}$. Hence, we get that

$$\begin{aligned} \left| \mathbb{E}[\exp(\beta_N(\omega - \mu) \mathbb{1}_{\{\omega \leq k_N\}})] - 1 - \sum_{i=1}^p \frac{\beta_N^i}{i!} \mathbb{E}[(\omega - \mu)^i] \right| \\ \leq \sum_{i=1}^p \frac{\beta_N^i}{i!} \mathbb{E}[|\omega - \mu|^i \mathbb{1}_{\{\omega > k_N\}}] + e^{\beta_N k_N} \beta_N^{p+1} \mathbb{E}[|\omega - \mu|^{p+1} \mathbb{1}_{\{\omega \leq k_N\}}]. \end{aligned}$$

And since $p < \alpha$, we have for any $i \leq p$ $\mathbb{E}[|\omega - \mu|^i \mathbb{1}_{\{\omega > k_N\}}] \leq C k_N^{i-\alpha}$. On the other hand, if $\alpha > p+1$, then $\mathbb{E}[|\omega - \mu|^{p+1} \mathbb{1}_{\{\omega \leq k_N\}}] \leq C$; if $\alpha < p+1$, then $\mathbb{E}[|\omega - \mu|^{p+1} \mathbb{1}_{\{\omega \leq k_N\}}] \leq C k_N^{p+1-\alpha}$; if $\alpha = p+1$, then $\mathbb{E}[|\omega - \mu|^{p+1} \mathbb{1}_{\{\omega \leq k_N\}}] \leq C \log k_N$. We therefore get that

$$\begin{aligned} \left| \mathbb{E}[\exp(\beta_N(\omega - \mu) \mathbb{1}_{\{\omega \leq k_N\}})] - 1 - \sum_{i=1}^p \frac{\beta_N^i}{i!} \mathbb{E}[(\omega - \mu)^i] \right| \\ \leq C(e^{\beta_N k_N} - 1) k_N^{-\alpha} + C e^{\beta_N k_N} (\beta_N k_N)^{p+1} \begin{cases} k_N^{-\alpha} & \text{if } p+1 > \alpha, \\ k_N^{-\alpha} \log k_N & \text{if } p+1 = \alpha, \\ k_N^{-(p+1)} & \text{if } p+1 < \alpha. \end{cases} \end{aligned}$$

Then, in the case where $\beta_N k_N \geq 1$, bounding $k_N^{p+1-\alpha}$ by $\beta_N^{\alpha-p-1}$ in the case $p+1 > \alpha$, we get (C.1). On the other hand, if $\beta_N k_N \leq 1$, considering only the case $\alpha < p+1$, we get (C.2). \square

References

- [1] T. Alberts, K. Khanin, and J. Quastel. The intermediate disorder regime for directed polymers in dimension $1 + 1$. *Ann. Probab.*, 42:1212–1256, 2014. [MR3189070](#)
- [2] A. Auffinger and O. Louidor. Directed polymers in random environment with heavy tails. *Comm. on Pure and Applied Math.*, 64:183–204, 2011. [MR2766526](#)
- [3] I. Benjamini and D. Wilson. Excited random walk. *Electron. Commun. Probab.*, 8:86–92, 2003. [MR1987097](#)
- [4] N. Berestycki and R. Cerf. The random walk penalised by its range in dimensions $d \geq 3$. *Annales Henri Lebesgue*, 4:1–79, 2021. [MR4213155](#)
- [5] Q. Berger, C.-H. Huang, N. Torri, and R. Wei. One-dimensional polymers in random environments: stretching vs. folding. [arXiv:2002.06899](#), 2020.
- [6] Q. Berger and H. Lacoin. The scaling limit of the directed polymer with power-law tail disorder. *Communications in Mathematical Physics*, pages 1–55, 2021. [MR4294286](#)
- [7] Q. Berger and N. Torri. Directed polymers in heavy-tail random environment. *Ann. Probab.*, 47(6):4024–4076, 2019. [MR4038048](#)
- [8] Q. Berger and N. Torri. Entropy-controlled last-passage percolation. *Ann. Appl. Probab.*, 29(3):1878–1903, 2019. [MR3914559](#)
- [9] Q. Berger and N. Torri. Beyondammersley’s last-passage percolation: a discussion on possible local and global constraints. *Annales de l’Institut Henri Poincaré D*, 8(2):213–241, 2021. [MR4261671](#)
- [10] G. Biroli, J. Bouchaud, and M. Potters. Extreme value problems in random matrix theory and other disordered systems. *J. Stat. Mech.*, 7:P07019, 15 pp., 2007. [MR2335697](#)
- [11] E. Bolthausen. Localization of a two-dimensional random walk with an attractive path interaction. *Ann. Probab.*, pages 875–918, 1994. [MR1288136](#)
- [12] N. Bouchot. One-dimensional polymer in a repulsive, random environment. *Work in progress*, 2022+.
- [13] F. Caravenna, P. Carmona, and N. Pétrelis. The discrete-time parabolic Anderson model with heavy-tailed potential. *Ann. Inst. Henri Poincaré: Probab. Stat.*, 48(4):1049–1080, 2012. [MR3052403](#)
- [14] F. Caravenna, R. Sun, and N. Zygouras. The continuum disordered pinning model. *Probab. Theory Related Fields*, 164:17–59, 2016. [MR3449385](#)
- [15] F. Caravenna, R. Sun, and N. Zygouras. Polynomial chaos and scaling limits of disordered systems. *J. Eur. Math. Soc. (JEMS)*, 19:1–65, 2017. [MR3584558](#)
- [16] X. Chen. *Random walk intersections : large deviations and related topics*. Mathematical surveys and monographs. American Mathematical Society, 2009. [MR2584458](#)
- [17] F. Comets. *Directed Polymers in Random Environments: École d’Été de Probabilités de Saint-Flour XLVI —2016*, volume 2175. Springer, 2017. [MR3444835](#)
- [18] H. David. *Order Statistics*. 2nd ed. Wiley Series in Probability and Mathematical Statistics. Wiley, New York, 1981. [MR0597893](#)
- [19] P. S. Dey and N. Zygouras. High temperature limits for $(1 + 1)$ -dimensional directed polymer with heavy-tailed disorder. *Ann. Probab.*, 44(6):4006–4048, 2016. [MR3572330](#)
- [20] J. Ding, R. Fukushima, R. Sun, and C. Xu. Geometry of the random walk range conditioned on survival among bernoulli obstacles. *Probabi. Theory Relat. Fields*, 177(1):91–145, 2020. [MR4095012](#)
- [21] M. Donsker and S. Varadhan. On the number of distinct sites visited by a random walk. *Commun. Pure Appl. Math.*, 32(6):721–747, 1979. [MR0539157](#)
- [22] W. Feller. *An introduction to probability theory and its applications*, volume 1. Wiley & Sons, 1968. [MR0228020](#)
- [23] T. Gueudre, P. Le Doussal, J.-P. Bouchaud, and A. Rosso. Ground-state statistics of directed polymers with heavy-tailed disorder. *Phys. Rev. E*, 91:062110, Jun 2015. [MR3491341](#)
- [24] B. Hambly and J. B. Martin. Heavy tails in last-passage percolation. *Probability Theory and Related Fields*, 137:227–275, 2007. [MR2278457](#)

- [25] C.-H. Huang. The scaling limits for wiener sausages in random environments. [arXiv:1902.04930](#), 2019.
- [26] N. C. Jain and W. E. Pruitt. The range of transient random walk. *Journal d'Analyse Mathématique*, 24(1):369–393, 1971. [MR0283890](#)
- [27] O. Kallenberg. *Foundations of Modern Probability*. Springer Probability and its Applications. Springer-Verlag New York, 2 edition, 2002. [MR1876169](#)
- [28] W. König, H. Lacoin, P. Mörters, and N. Sidorova. A two cities theorem for the parabolic Anderson model. *Ann. Probab.*, 37(1):347–392, 2009. [MR2489168](#)
- [29] M. Lässig and H. Kinzelbach. Upper critical dimension of the Kardar-Parisi-Zhang equation. *Physical Review Letters*, 78(5):903, 1997.
- [30] C. Stone. On local and ratio limit theorems. In *Proc. Fifth Berkeley symposium on mathematical statistics and probability*. Berkeley, pages 217–224, 1967. [MR0222939](#)
- [31] N. Torri. Pinning model with heavy tailed disorder. *Stochastic Processes and their Applications*, 126:542–571, 2016. [MR3434993](#)
- [32] K. Uchiyama. The first hitting time of a single point for random walks. *Electron. J. Probab.*, 16:1960–2000, 2011. [MR2851052](#)

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