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Bayesian mixture models (in)consistency for the number of clusters

Louise Alamichel¹, Daria Bystrova¹, Julyan Arbel¹, Guillaume Kon Kam King²

¹Univ. Grenoble Alpes, Inria, France ²Univ. Paris-Saclay, INRAE, France



Mixture models

$$x \sim \int p(\cdot|\boldsymbol{\theta})G(\mathrm{d}\boldsymbol{\theta}),$$

with G the mixing measure.

$$x \mid \boldsymbol{w}, \boldsymbol{\theta}, \boldsymbol{K} \sim f(\cdot) = \sum_{k=1}^{K} w_k p(\cdot | \theta_k),$$

where K is the number of components, w the weights and p a kernel parametrized by $\boldsymbol{\theta}$. K may be fixed, infinite or random.

Posterior consistency

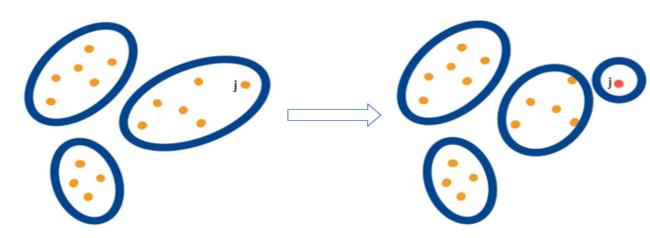
The posterior distribution is said to be consistent at θ_0 if $\Pi(U^c \mid X_{1:n}) \xrightarrow[n \to \infty]{} 0$ in P_{θ_0} -probability for all neighborhoods U of θ_0 .

 G_0 defines a finite mixture with K_0 components; consistency for G with K components:

Quantity	Finite		Infinite	MFM
	$K = K_0$	$K \ge K_0$	$K = \infty$	K random
Density f_0	√	√	√	√
G_0				
K_0	N/A	X/V	X/V	

Partition-based model

- Partitions in k sets:
- $\mathcal{A}_k(n) = \{(A_1, \dots, A_k) : \bigcup_{i=1}^k A_i = 1 : n\}$
- For $A \in \mathcal{A}_k(n)$, $\mathcal{Z}_A = \{B(A, j), j \in 1 : n\}$

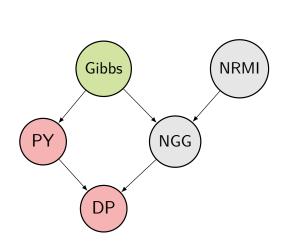


 $A \in \mathcal{A}_k(n)$

 $B = B(A, j) \in \mathcal{A}_{k+1}(n)$

- K_n : number of clusters
- $p(A, k) = p(A)\mathbb{I}(A \in \mathcal{A}_k(n))$ $\theta_i|A,k\sim\pi,$ $i=1,\ldots,n$ $x_j|A, k, \theta_{1:k} \sim p(\cdot|\theta_\ell) \text{ if } j \in A_\ell, \quad j = 1, \ldots, n$

Partition distributions

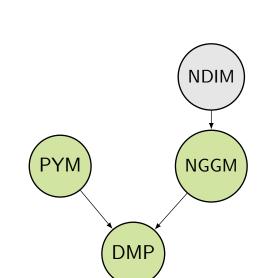


Gibbs-type process with $0 < \sigma < 1, A \in \mathcal{A}_k(n), n_j := |A_j|$: $p(A) = \frac{V_{n,k}}{k!} \prod_{j=1}^{k} (1 - \sigma)_{n_j - 1}.$

 $V_{n,k}$ satisfy the recurrence relation

 $V_{n,k} = (n - \sigma k)V_{n+1,k} + V_{n+1,k+1}, V_{1,1} = 1.$

Examples of Gibbs-type processes finite representation:



- Dirichlet multinomial process,
- Pitman-Yor multinomial process,
- NGG multinomial process,

parametrized by $K < \infty$: for $k \leq \min(n, K), A \in \mathcal{A}_k(n),$

$$p(A) = \Pi_{K}(n_1, \dots, n_k).$$

Inconsistency Theorem [1]

Condition 1: Assume for any $k \geq 1$, $\limsup_{n\to\infty} \frac{1}{n} \max_{A\in\mathcal{A}_k(n)} \max_{B\in\mathcal{Z}_A} \frac{p(A)}{p(B)} < \infty.$

Condition 2: Condition on the data distribution which involves a control on the likelihood.

Theorem 1 [1]: Let $X_1, X_2, \ldots \in \mathcal{X}$ be a sequence of r.v., and consider a partition-based model. Then, for any $k \geq 1$ if Conditions 1 and 2 hold, we have

 $\limsup \Pi(K_n = k | X_{1:n}) < 1$ with probability 1.

Inconsistency results

Proposition 1: Consider a Gibbs-type process with 0 < $\sigma < 1$, then Condition 1 holds for any $k \in \{1, 2, \ldots\}$, so does the inconsistency of Theorem 1.

Proposition 2: Consider any of the following priors:

- Dirichlet multinomial process,
- Pitman-Yor multinomial process,
- Normalized Generalized Gamma multinomial process, then Condition 1 holds for any $k < \min(n, K)$, and so does the inconsistency of Theorem 1.

Illustration on simulated data:

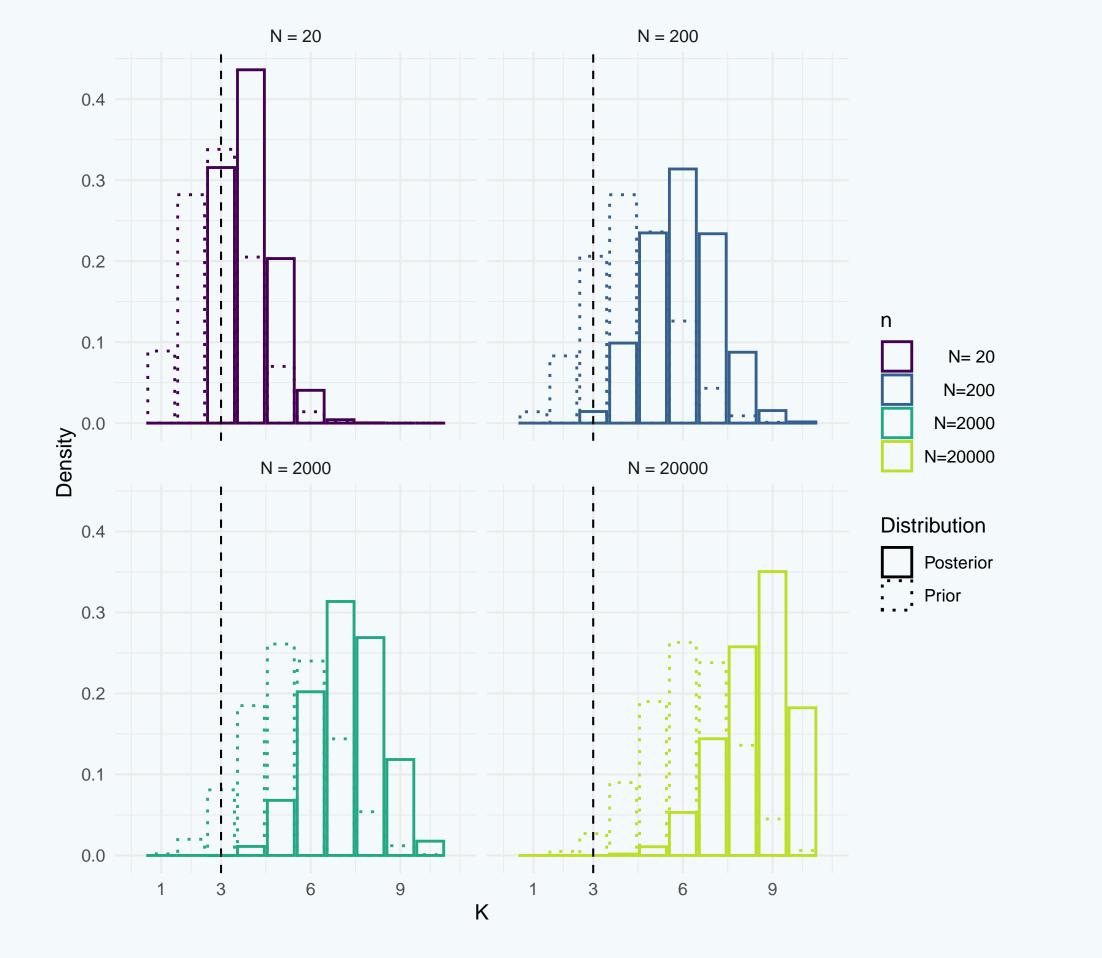
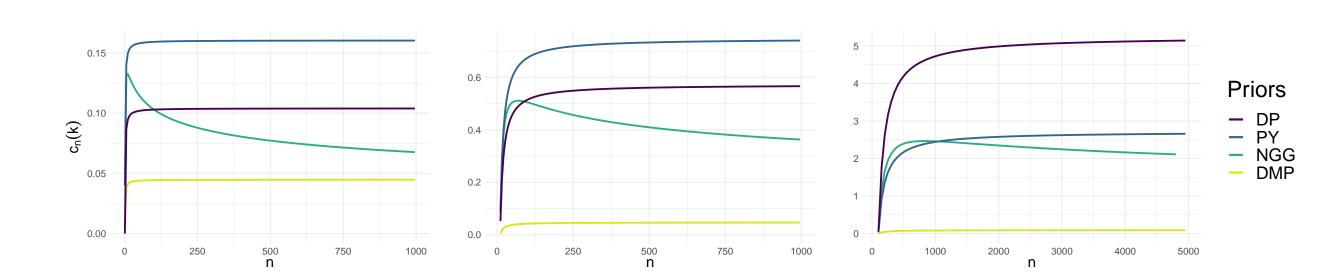


Figure 1: Distribution of K_n for a Dirichlet multinomial process, $\alpha = 0.9$.

Idea of the proof for Gibbs-type processes

- For fixed $k, \forall A \in \mathcal{A}_k(n)$ and B = B(A, j), we want $\frac{1}{n} \max_{A,B} \frac{p(A)}{p(B)} < \infty$, we can show $\frac{1}{n} \frac{p(A)}{p(B)} \le \frac{V_{n,k}}{V_{n,k+1}} (k+1)$.
- It is sufficient to prove $\left(\frac{V_{n,k}}{V_{n,k+1}}\right)_{n\geq 0}$ is bounded. By an approximation of $V_{n,k}$: $\left(\frac{V_{n,k}}{V_{n,k+1}}\right)_{n>0}$ converges, so is bounded.

Illustration of function $n \mapsto \frac{1}{n} \max_{A \in \mathcal{A}_k(n)} \max_{B \in \mathcal{Z}_A} \frac{p(A)}{n(B)}$, for $k \in \{1, 10, 100\}$:



Consistency results

- Dirichlet multinomial process: satisfies assumptions in [2] for some α , so the weights of extra clusters $\rightarrow 0$, a form of mixing measure consistency.
- PY multinomial process: for $\sigma = 1/2$ and some α satisfies assumptions in [2], the weights of extra clusters $\rightarrow 0$.
- Dirichlet process: Posterior consistency for K_n by MTM shown in [3].
- Pitman-Yor process: from [4] we deduce a contraction rate, hence by using MTM algorithm we have posterior consistency for K_n .
- Overfitted mixtures: Posterior consistency for K_n by MTM.

Merge-Truncate-Merge algorithm [3]

Suppose $\omega_n: W_r(G,G_0) = o_p(\omega_n)$ for posterior sample G

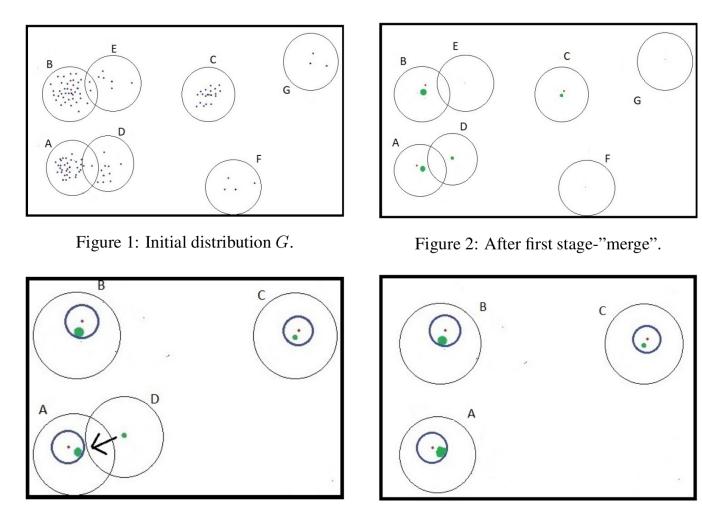


Figure 2: MTM algorithm illustration (figure from [3]).

Figure 4: After second stage-"merge".

k number of support atoms of the output G.

Figure 3: After second stage-"truncation"

 $\Pi\left(\tilde{k}=k_0\mid X_{1:n}\right)\to 1$ in probability.

Future work

- Do the inconsistency results generalize to NIDM processes?
- Can consistency be recovered by adding a prior on α , as in [5]?

Preprint available on arXiv, search for "Alamichel".

References

- [1] Jeffrey W Miller and Matthew T Harrison. Inconsistency of Pitman-Yor process mixtures for the number of components. The Journal of Machine Learning Research, 15(1):3333–3370, 2014.
- [2] Judith Rousseau and Kerrie Mengersen. Asymptotic behaviour of the posterior distribution in overfitted mixture models. Journal of the Royal Statistical Society: Series B (Statistical Methodology), 73(5):689–710, 2011.
- [3] Aritra Guha, Nhat Ho, and XuanLong Nguyen. On posterior contraction of parameters and interpretability in bayesian mixture modeling. Bernoulli, 27(4):2159–2188, 2021.
- [4] Catia Scricciolo. Adaptive Bayesian density estimation in Lp-metrics with Pitman-Yor or normalized inverse-Gaussian process kernel mixtures. Bayesian Analysis, 9(2):475–520, 2014.
- [5] Filippo Ascolani, Antonio Lijoi, Giovanni Rebaudo, and Giacomo Zanella. Clustering consistency with Dirichlet process mixtures. arXiv preprint arXiv:2205.12924, 2022.