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### **To cite this version:**

Clément Foucart, Martin Möhle. Asymptotic behaviour of ancestral lineages in subcritical continuousstate branching populations. Stochastic Processes and their Applications, 2022, 150, pp.510-531.  $10.1016/j.\text{spa}.2022.05.001$ . hal-03860204

## **HAL Id: hal-03860204 <https://hal.science/hal-03860204v1>**

Submitted on 22 Jul 2024

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stochastic processes and their applications

[Stochastic Processes and their Applications 150 \(2022\) 510–531](https://doi.org/10.1016/j.spa.2022.05.001)

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### Asymptotic behaviour of ancestral lineages in subcritical continuous-state branching populations

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Received 7 December 2020; received in revised form 25 November 2021; accepted 4 May 2022 Available online 12 May 2022

#### Abstract

Consider the population model with infinite size associated to subcritical continuous-state branching processes (CSBP). We study the flow of ancestral lineages as time goes to the past and show that, properly renormalized, it converges almost surely to the inverse of a drift-free subordinator whose Laplace exponent is explicit in terms of the branching mechanism. The inverse subordinator is shown to be partitioning the current population into ancestral families with distinct common ancestors. When Grey's condition is satisfied, the population comes from a discrete set of ancestors and the ancestral families have i.i.d. sizes distributed according to the quasi-stationary distribution of the CSBP conditioned on non-extinction. When Grey's condition is not satisfied, the population comes from a continuum of ancestors which is described as the set of increase points  $\mathscr S$  of the limiting inverse subordinator. The proof is based on a general result for stochastically monotone processes of independent interest, which relates θ-invariant measures and θ-invariant functions for a process and its Siegmund dual. © 2022 Elsevier B.V. All rights reserved.

*MSC:* 60J80; 60J70; 92D25

*Keywords:* Branching processes; Continuous-state space; Inverse subordinators; Ancestral lineage; Siegmund dual; Invariant function

#### 1. Introduction

Continuous-state branching processes (CSBPs) are positive Markov processes satisfying the branching property. They arise as scaling limits of Galton–Watson processes and form

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<https://doi.org/10.1016/j.spa.2022.05.001>

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a fundamental class of random population models. Their longterm behaviour has received a great deal of attention since the seventies. We refer for instance to the early works of Bingham [6] and Grey [16]. In the seminal work [5], Bertoin and Le Gall showed how to encode a complete genealogy of a random branching population by considering a flow of subordinators  $(X_s, t(x)) - \infty < s < t < \infty, x > 0$ ). In this setting, for any  $x \in (0, \infty)$ , the process  $(X_t(x), t \ge 0) := (X_{0,t}(x), t \ge 0)$  is a CSBP started from *x*, viewed as the size of the progeny of all individuals lying in the interval  $(0, x)$  at time 0.

The initial value  $x$  being arbitrarily large, the population in the flow of subordinators representation has an infinite size at all times. The genealogy is defined as follows: individuals at time *s* with descendants at time *t* > *s* are the locations of the jumps of the subordinator  $x \mapsto X_{s,t}(x)$  and the descendants at time *t* of the individuals in the population at time *s* are represented by the jump intervals. The flow being defined on  $\mathbb{R}$ , all individuals have arbitrarily old ancestors. We shall provide more background on the flow of subordinators in the sequel.

Most works on CSBPs focus on their long-term behaviour forward in time. We refer to Bertoin et al. [4], Duquesne and Labbé [11], Labbé [19] and Foucart and Ma [13] for studies in the framework of the flow of subordinators. In this article, we are interested in the backward genealogy of the continuous population and how it behaves on the long-term. To the best of our knowledge, fewer works on CSBPs have been done in this direction. We refer however to the works of Lambert [21], Lambert and Popovic [23], Johnston and Lambert [18] and Foucart et al. [14]. The representation of the population model through  $(X_{s,t}(x), s \le t, x \ge 0)$  allows one to follow the ancestral lineages *backward in time*. The work [14] initiates the study of the inverse flow  $(\hat{X}_{s,t}(x), s \le t, x \ge 0)$  defined for  $s \le t$  and  $x \in (0, \infty)$ , as

$$
\hat{X}_{s,t}(x) := \inf \{ y \ge 0 : \ X_{-t,-s}(y) > x \}. \tag{1.1}
$$

This random variable represents the ancestor at time  $-t$  of the individual  $x$  in the population at time −*s*. From now on, we consider an arrow of time pointing to the past, and call  $\hat{X}_t(x) := \hat{X}_{0,t}(x)$ , the ancestor at time  $t \ge 0$  (backwards) of the individual *x* of the population at time 0. Notice that it can also be interpreted as the minimal size which the population at time *t* (backwards) should have for the current population to have a size greater than *x*. The two-parameter flow  $(\hat{X}_t(x), t \ge 0, x \ge 0)$  is representing the ancestral lineages of the individuals in the current population. We call  $(\hat{X}_t(x), t \ge 0)$  the *ancestral lineage process*. This is a Feller process with no positive jumps. Moreover, for any  $x \neq y$ , whenever  $(\hat{X}_t(x), t \geq 0)$ and  $(\hat{X}_t(y), t \ge 0)$  meet, they coalesce and such a coalescence represents the occurrence in the past of a common ancestor of the individuals  $x$  and  $y$ . We refer to  $[14]$  for a study of the coalescent processes embedded in the flow  $(\hat{X}_t(x), t \ge 0, x \ge 0)$ .

For any fixed  $x \in (0, \infty)$ , the Markov processes  $(\hat{X}_t(x), t \ge 0)$  and  $(X_t(x), t \ge 0)$  are linked through a duality relationship, called *Siegmund duality*, of the following form: for any  $t \geq 0$ and *x*,  $y \in (0, \infty)$ 

$$
\mathbb{P}(\hat{X}_t(x) < y) = \mathbb{P}(x < X_t(y)).\tag{1.2}
$$

We refer to Siegmund [29] and Clifford and Sudbury [9] for a general study of the duality (1.2). See [14, Equation (3.5), Section 3] for the case of CSBPs.

We wish to mention that Siegmund dual processes of discrete branching Markov processes have been already studied by Asmussen and Sigman [1, Example 9], Li et al. [27] and more recently by Pakes [28]. However no mention was made in these works about the genealogical interpretation of the Siegmund duals and their almost sure renormalization. Besides the fact that our study lies in the continuous state space setting, one of the main contributions of this article will be to provide a genealogical interpretation of the Siegmund duals and their limits.

We focus in this article on subcritical CSBPs. In such a setting, it has been shown in [14] that for all  $x \in (0, \infty)$ , the Markov process  $(\hat{X}_t(x), t \ge 0)$  is transient. The main aim of this article is to obtain an almost sure renormalization of the inverse flow  $\hat{X}$  of all subcritical CSBPs, including those with no quasi-stationary distributions, namely those for which Grey's condition does not hold, see Section 2. The limit process is the inverse of a certain subordinator with an explicit Laplace exponent. Last but not least, we establish that this inverse subordinator is partitioning the current population into ancestral families with distinct common ancestors.

The paper is organized as follows. Further background on CSBPs and their representation in terms of flow of subordinators are provided in Section 2. Fundamental properties of the ancestral lineage process  $(\hat{X}_t(x), t \ge 0)$ , such as its Siegmund duality relation with the CSBP  $(X_t(x), t \ge 0)$  and the representation of its semigroup, are also recalled. Our main results are stated in Section 3 and proven in Section 4. The proof is based on a general result, established in Theorem 4.1, for stochastically monotone Markov processes by showing how to link (infinite) θ-invariant measures of a process (*Xt*, *t* ≥ 0) with (increasing) θ-invariant functions of its Siegmund dual process  $(\hat{X}_t, t \ge 0)$ . We apply this result in the setting of CSBPs.

#### 2. Background on CSBPs and the flow of subordinators

We first recall basic definitions and properties of CSBPs and their representation in terms of flows. These processes are continuous time and continuous space analogue of Galton–Watson Markov chains. They have been introduced by Lamperti  $[24]$  and Jiřina  $[17]$ . CSBPs are positive Markov processes satisfying the branching property: for any  $x, y \ge 0$  and fixed time  $t \ge 0$ ,

$$
X_t(x + y) = X'_t(x) + X''_t(y),
$$
\n(2.3)

where  $(X_t(x + y), t \ge 0)$  is a CSBP started from  $x + y$ , and  $(X'_t(x), t \ge 0)$  and  $(X''_t(y), t \ge 0)$ are two independent copies of the process started respectively from *x* and *y*. We refer the reader to [26, Chapter 3] for an introduction to CSBPs. Denote by  $\mathcal L$  the generator of  $(X_t(x), t \ge 0)$ . For any  $q \ge 0$ , set  $e_q(x) := e^{-qx}$  for any  $x \ge 0$ . The operator  $\mathcal L$  acts on the exponential functions as follows. For all  $q, x \geq 0$ ,

$$
\mathcal{L}e_q(x) = \Psi(q)x e_q(x),\tag{2.4}
$$

where  $\Psi$  is a Lévy–Khintchine function and is called the branching mechanism. We refer e.g. to Silverstein [30]. The linear span of exponential functions  $A := Span({e_q, q \in [0, \infty)}$  is a core for generator  $\mathcal{L}$ .

We shall merely be interested in subcritical CSBPs for which  $\Psi$  is of the form

$$
\Psi(u) = \frac{\sigma^2}{2}u^2 + \gamma u + \int_0^\infty (e^{-ux} - 1 + ux)\pi(\mathrm{d}x) \text{ for all } u \ge 0,
$$
\n(2.5)

where  $\gamma = \Psi'(0+) > 0$ ,  $\sigma \ge 0$ , and  $\pi$  is a Lévy measure, i.e. a Borel measure such that  $\int_0^\infty (x \wedge x^2) \pi(dx) < \infty$ . We assume that either  $\pi \neq 0$  or  $\sigma > 0$ , so that  $\Psi$  is not linear. The semigroup of  $(X_t(x), t \ge 0)$  satisfies for any  $\lambda \in (0, \infty), t \ge 0$  and  $x \in [0, \infty)$ 

$$
\mathbb{E}[e^{-\lambda X_t(x)}] = e^{-xv_t(\lambda)},\tag{2.6}
$$

with for any  $\lambda \in (0, \infty)$ ,  $t \mapsto v_t(\lambda)$  defined as the solution to the integral equation

$$
\int_{v_t(\lambda)}^{\lambda} \frac{\mathrm{d}u}{\Psi(u)} = t. \tag{2.7}
$$

Note that  $t \mapsto v_t(\lambda)$  solves  $\frac{d}{dt}v_t(\lambda) = -\Psi(v_t(\lambda))$  with  $v_0(\lambda) = \lambda$ . As a first consequence of (2.6), the process  $(X_t(x), t \ge 0)$  is extinct at time *t* with probability  $e^{-xv_t(\infty)}$ , where  $v_t(\infty) := \lim_{\lambda \to \infty} v_t(\lambda) \in (0, \infty]$ . The latter is finite if and only if  $\Psi$  satisfies Grey's condition

$$
\int^{\infty} \frac{\mathrm{d}u}{\Psi(u)} < \infty. \tag{2.8}
$$

Lambert [22] and Li [25] have studied the subcritical CSBP ( $X_t(x)$ ,  $t \ge 0$ ) conditioned on non-extinction and established the following weak convergence when (2.8) holds:

$$
\mathbb{P}(X_t(x) \in \cdot \mid X_t(x) > 0) \xrightarrow[t \to \infty]{} \nu_\infty(\cdot),
$$

where  $v_{\infty}$  is the (minimal) quasi-stationary distribution of the CSBP and has Laplace transform

$$
\int_0^\infty e^{-qx} \nu_\infty(\mathrm{d}x) = 1 - \kappa_\infty(q) := 1 - e^{-\Psi'(0) + \int_q^\infty \frac{\mathrm{d}u}{\Psi(u)}}, \ q \ge 0. \tag{2.9}
$$

An interesting phenomenon for CSBPs is that when Grey's condition (2.8) does not hold, the latter are persistent in the sense that although subcritical, they are not getting absorbed at the boundary 0, but are decreasing towards 0 while keeping positive mass at all times. In particular, no quasi-stationary distributions exist in this setting.

The branching property  $(2.3)$  can be translated in terms of independence and stationarity of the increments of a process  $(X_t(x), x \ge 0)$  for any fixed time *t*. The latter is therefore a subordinator and according to (2.6), its Laplace exponent is  $\lambda \mapsto v_t(\lambda)$ . Starting from this observation, Bertoin and Le Gall in [5] showed that a complete population model can be associated to CSBPs through a flow of subordinators  $(X_{s,t}(x), s \le t, x \ge 0)$ .

More precisely, the collection of processes  $(X_{s,t}(x), s \le t, x \ge 0)$  is satisfying the following properties:

- (1) For every  $s \leq t$ ,  $x \mapsto X_{s,t}(x)$  is a càdlàg subordinator with Laplace exponent  $\lambda \mapsto$  $v_{t-s}(\lambda)$ .
- (2) For every  $t \in \mathbb{R}$ ,  $(X_{r,s}, r \leq s \leq t)$  and  $(X_{r,s}, t \leq r \leq s)$  are independent.
- (3) For every  $r \le s \le t$ ,  $X_{r,t} = X_{s,t} \circ X_{r,s}$  a.s.

The two-parameter flow  $(X_t(x), x \geq 0, t \geq 0) := (X_{0,t}(x), x \geq 0, t \geq 0)$  is a flow of CSBPs with branching mechanism  $\Psi$ , each starting from an initial population of arbitrarily large size *x*. The three-parameter flow above provides a complete genealogy of the underlying (infinite) population: let  $y \in (0, \infty)$ , if  $X_{s,t}(y-) < X_{s,t}(y)$ , then the individual y at time *s* has descendants at time *t* and those are represented by the interval  $(X_{s,t}(y-), X_{s,t}(y))$ ; see Fig. 1.

We will now always work on the probability space on which the flow of subordinators  $(X_{s,t}(x), s \le t, x \ge 0)$  and thus the inverse flow  $(\hat{X}_{s,t}(x), t \ge s, x \ge 0)$ , see (1.1), are defined. We stress that the statements below hold true with general branching mechanisms, including those that are not subcritical. We shall explicitly mentioned the condition  $\Psi'(0+) > 0$  when this is needed. Similarly as for the forward flow, we summarize here fundamental properties of the inverse flow, see [14, Section 3]

- (1) For every  $t \geq s$ ,  $x \mapsto \hat{X}_{s,t}(x)$  is a càdlàg inverse subordinator.
- (2) For every  $t \in \mathbb{R}$ ,  $(\hat{X}_{r,s}, r \leq s \leq t)$  and  $(\hat{X}_{r,s}, t \leq r \leq s)$  are independent.
- (3) For all  $r \leq s \leq t$ ,  $\hat{X}_{r,t} = \hat{X}_{s,t} \circ \hat{X}_{s,r}$  a.s.

The ancestral lineage process, defined as

 $(\hat{X}_t(x), t \ge 0) := (\hat{X}_{0,t}(x), t \ge 0),$ 



Fig. 1. Schematic representation of the genealogy forward in time.

is a non-explosive càdlàg Feller process with no positive jumps, namely for all  $x \in (0, \infty)$ ,  $\hat{T}_{\infty} := \inf\{t > 0 : \hat{X}_t(x) = \infty\} = \infty$  a.s. and  $\mathbb{P}(\sup_{t > 0} (\hat{X}_t(x) - \hat{X}_{t-}(x)) > 0) = 0$ . Property (3) entails that the flow of processes  $(\hat{X}_t(x), t \geq 0, x \in (0, \infty))$  is *coalescing*, in the sense that, when two ancestral lineages  $(\hat{X}_t(x), t \ge 0)$  and  $(\hat{X}_t(y), t \ge 0)$  meet, they merge. Such a coalescence represents the occurrence in the past of a common ancestor of the individuals *x* and *y*. We refer the reader to [14, Section 5.2] for a study of coalescent processes embedded in the flow  $(\hat{X}_{s,t}, s \le t)$ . The next theorem characterizes the semigroup of  $(\hat{X}_t, t \ge 0)$ . We denote it by  $(Q_t, t \geq 0)$ .

Theorem 2.1 (*Theorem 3.5, Proposition 3.6 in [14]*). *For any continuous function f defined on*  $(0, \infty)$  *and any*  $q > 0$ *,* 

$$
\mathbb{E}[Q_t f(\mathbf{e}_q)] = \mathbb{E}[f(\mathbf{e}_{v_t(q)})],\tag{2.10}
$$

*where for any*  $\lambda \in (0, \infty)$ ,  $e_{\lambda}$  *is an exponential random variable with parameter*  $\lambda$ *. The process*  $\hat{X}$  admits an entrance boundary at  $0+$  *if and only if* (2.8) *is satisfied.* 

When (2.8) holds, the process started from 0,  $(\hat{X}_t(0+), t \ge 0)$ , is defined at any time as  $\hat{X}_t(0+) := \lim_{x \downarrow 0} \hat{X}_t(x)$  a.s. This corresponds to the first individual at generation *t* backwards in time, with descendants at time 0.

We now verify that  $\hat{X}$  satisfies some properties of regularity. For *x*,  $y \in (0, \infty)$ , set

$$
\hat{T}_y := \inf\{t > 0 : \hat{X}_t(x) > y\} = \inf\{t > 0 : \hat{X}_t(x) = y\}.
$$
\n(2.11)

We shall sometimes write  $\hat{T}_y = \hat{T}_y^x$  to emphasize on the initial state *x* of the process.

**Lemma 2.2** (*Regularity*). *If*  $-\Psi$  *is not the Laplace exponent of a subordinator, i.e.*  $\Psi(u) \ge 0$  *for some u* > 0*, then the process*  $\hat{X}$  *is regular on*  $(0, \infty)$ *, namely for any*  $x < y$ *,*  $\mathbb{P}_x(\hat{T}_y < \infty) > 0$ *.* 

**Proof.** Let  $x, y \in (0, \infty)$ , for any  $t > 0$ ,  $\mathbb{P}_x(\hat{T}_y < t) \geq \mathbb{P}(\hat{X}_t(x) > y) = \mathbb{P}(x > X_t(y))$ . By assumption,  $-\Psi$  is not the Laplace exponent of a subordinator, this ensures that the CSBP is not almost surely non-decreasing, see e.g. [6, page 220], and that the event  $\{X_t(y) \longrightarrow_{t \to \infty} 0\}$  has positive probability. Hence,  $\mathbb{P}_x(\hat{T}_y < \infty) \ge \lim_{t \to \infty} \mathbb{P}(x > X_t(y)) \ge \mathbb{P}(X_t(y) \to 0) > 0.$  □

In the subcritical case, for which  $\Psi'(0+) = \gamma > 0$ , one has  $\lim_{t\to\infty} v_t(\lambda) = 0$  and all families forward in time are getting extinct. As mentioned in the introduction, in this case the ancestral lineages are transient.

**Proposition 2.3** (*Proposition 3.8 in [14]*). *Assume*  $\Psi'(0+) > 0$ *. For any*  $x \in (0, \infty)$ *, the ancestral lineage process*  $(\hat{X}_t(x), t \ge 0)$  *is transient, i.e.*  $\hat{X}_t(x) \longrightarrow_{t \to \infty} \infty$  *a.s.* 

The following versions of Siegmund duality relationships will be more convenient to work with in the sequel. They will allow us to apply a general result, Theorem 4.1, about invariant functions and invariant measures for a given process and its Siegmund dual.

**Lemma 2.4.** *For any t*  $\geq 0$  *and x*,  $y \in (0, \infty)$ *,* 

$$
\{\hat{X}_t(x) > y\} = \{x > X_{-t,0}(y)\} \text{ almost surely,}
$$
\n(2.12)

*and*

$$
\mathbb{P}(\hat{X}_t(x) > y) = \mathbb{P}(x > X_t(y)).\tag{2.13}
$$

**Remark 2.5.** We stress here on the strict inequalities in  $(2.13)$ . This form of duality is actually the one used by Siegmund in his fundamental article [29].

**Proof.** Let  $t \ge 0$  and  $x, y \in (0, \infty)$ . We establish that  $\{\hat{X}_t(x) \le y\} = \{x \le X_t(y)\}\$ almost surely. Note that this is equivalent to  $(2.12)$ . According to [14, Lemma 3.3-(1), Section 3],  ${\{\hat{X}_t(x) \leq y\}} = {x < X_t(y)}$  almost surely, therefore we only need to focus on the events  $\{\hat{X}_t(x) = y\}$  and  $\{X_t(y) = x\}$ . Recall the definition of  $\hat{X}_t(x)$  in (1.1),

$$
\{\hat{X}_t(x) = y\} = \{X_{-t,0}(y-) \le x < X_{-t,0}(y)\} \cup \{X_{-t,0}(y-) = X_{-t,0}(y) = x\}.
$$

Since  $X_{-t,0}$  is a subordinator, it has no almost sure fixed discontinuities and the event  ${X_{-t,0}(y-) < X_{-t,0}(y)}$  for fixed *y* has probability 0. Thus

 $\{\hat{X}_t(x) = y\} = \{X_{-t,0}(y) = x\}$  a.s.,

and (2.12) is established. Since  $X_{-t,0}(y)$  has the same law as  $X_t(y)$ ,

$$
\mathbb{P}(\hat{X}_t(x) = y) = \mathbb{P}(X_{-t,0}(y) = x) = \mathbb{P}(X_t(y) = x),
$$

and the identity  $(2.13)$  holds.  $\Box$ 

#### 3. Results

Assume that the branching mechanism  $\Psi$  is subcritical, namely  $\Psi'(0+) > 0$ . For any  $\lambda \in (0, \infty)$ , define the map on  $[0, \infty)$ :

$$
\kappa_{\lambda} : \theta \mapsto e^{-\Psi'(0+) \int_{\theta}^{\lambda} \frac{\mathrm{d}u}{\Psi(u)}}. \tag{3.14}
$$

We shall see that  $\kappa_{\lambda}$  is the Laplace exponent of a drift-free subordinator, namely it takes the form  $\kappa_{\lambda}(\theta) = \int_0^{\infty} (1 - e^{-\theta x}) \nu_{\lambda}(\mathrm{d}x)$ , for any  $\theta \ge 0$ , where  $\nu_{\lambda}$  is a Lévy measure, i.e. a Borel measure on  $(0, \infty)$  such that  $\int_0^\infty (1 \wedge x) v_\lambda(dx) < \infty$ . The latter is finite when  $\kappa_\lambda(\infty) < \infty$ . This occurs if and only if Grey's condition holds,  $\int_{\infty}^{\infty} \frac{du}{\psi(u)} < \infty$ , see (2.8). In this case, the function  $\kappa_{\infty}$  defined in (2.9), is the Laplace exponent of a compound Poisson process with jump law  $v_{\infty}$ , the quasi-stationary distribution of the  $\Psi$ -CSBP conditioned on non-extinction.

**Theorem 3.1.** *Assume*  $\Psi'(0+) > 0$ *. Fix*  $\lambda \in (0, \infty)$ *. Then, almost surely* 

$$
v_t(\lambda)\hat{X}_t(x) \underset{t \to \infty}{\longrightarrow} \hat{W}^{\lambda}(x), \text{ for all } x \notin J^{\lambda} := \{x > 0 : \hat{W}^{\lambda}(x) > \hat{W}^{\lambda}(x-\),
$$

where the process  $\hat{W}^{\lambda}$  has càdlàg paths and its right-inverse process  $W^{\lambda}$ , defined for any  $y \geq 0$ *by*

$$
W^{\lambda}(y) := \inf\{x \ge 0 : \hat{W}^{\lambda}(x) > y\},\
$$

*is a drift-free subordinator with Laplace exponent*  $\kappa_{\lambda}$  *defined in* (3.14)*. In addition, the following dichotomy holds:*

- (i) *If*  $\int^{\infty} \frac{du}{\Psi(u)} < \infty$ , then for any  $\lambda \in (0, \infty]$  the process  $(\hat{W}^{\lambda}(x), x \ge 0)$  has piecewise *constant sample paths almost surely.*
- (ii) *If*  $\int^{\infty} \frac{du}{\Psi(u)} = \infty$ , then for any  $\lambda \in (0, \infty)$  the process  $(\hat{W}^{\lambda}(x), x \ge 0)$  has continuous *sample paths almost surely.*

**Remark 3.2.** In case (i), since the Lévy measure  $v_{\lambda}$  is finite,  $(\hat{W}^{\lambda}(x), x \ge 0)$  is the inverse of a compound Poisson process with Laplace exponent  $\kappa_{\lambda}$ . It is natural to choose  $\lambda = \infty$  in which case the following almost-sure convergence holds

$$
v_t(\infty)\hat{X}_t(x) \longrightarrow_{t \to \infty} \hat{W}^{\infty}(x) \text{ for all } x \notin J^{\infty},
$$

where  $(\hat{W}^{\infty}(x), x > 0)$  is the inverse of a compound Poisson process with Laplace exponent  $\kappa_{\infty}$ , whose Lévy measure is the quasi-stationary law of the CSBP. In case (ii), one cannot take  $\lambda = \infty$ , since when Grey's condition does not hold,  $v_t(\infty) = \infty$  for all  $t \ge 0$ .

Recall  $\pi$  the Lévy measure in the Lévy–Khintchine form (2.5) of  $\Psi$ . The following corollary shows that the ancestral lineage process  $(\hat{X}_t(x), t \ge 0)$  has an exponential growth when the measure  $\pi$  satisfies an  $L \log L$  condition.

**Corollary 3.3.** *For any*  $\lambda > 0$ ,  $v_t(\lambda) \sim c_{\lambda} e^{-\Psi'(0+)t}$  *for some constant*  $c_{\lambda} > 0$  *if and only if*  $\int_1^\infty u \log u \pi(\mathrm{d}u) < \infty$ *. Moreover, under this latter condition, almost surely* 

$$
e^{-\Psi'(0+)t}\hat{X}_t(x) \xrightarrow[t \to \infty]{} \hat{W}(x), \text{ for all } x \notin J := \{x > 0 : \hat{W}(x) > \hat{W}(x-\},
$$

*where*  $\hat{W}$  *is the inverse of a subordinator*  $W$  *with Laplace exponent* 

$$
\kappa: \theta \in [0, \infty) \mapsto \theta e^{-\Psi'(0+)\int_0^{\theta} \left(\frac{1}{\Psi'(0+)\mu} - \frac{1}{\Psi(u)}\right) du}.
$$

Remark 3.4. Corollary 3.3 is reminiscent to results for supercritical CSBPs with an *L* log *L* moment, which have an exponential growth when they are not getting extinct, see [16] and e.g. [26, Chapter 3].

**Example 3.5.** Let  $\gamma > 0$ . Consider the subcritical Neveu CSBP whose branching mechanism is defined by  $\Psi(u) := \gamma(u+1)\log(u+1)$  for all  $u \ge 0$ . Note that  $\Psi'(0+) = \gamma > 0$  and  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} = \infty$ . Solving (2.7) yields  $v_t(\lambda) = (\lambda + 1)^{e^{-\gamma t}} - 1 \sim_{t \to \infty} \log(1 + \lambda)e^{-\gamma t}$ , where  $\hat{W}$  is the inverse of a subordinator *W* with Laplace exponent

$$
\kappa(\theta) = \gamma \log(1+\theta) = \int_0^\infty (1 - e^{-\theta x}) \gamma \frac{e^{-x}}{x} dx.
$$

The limiting process  $\hat{W}$  is therefore an inverse Gamma subordinator.

The next observation ensures that the choice of the parameter  $\lambda$  is arbitrary in both cases (i) and (ii) in the sense that a change in  $\lambda$  only affects the limit by a multiplicative factor.

**Lemma 3.6.** *For any*  $\lambda' \neq \lambda \in (0, \infty)$  *and*  $x \in (0, \infty)$ *,* 

 $\hat{W}^{\lambda'}(x) = c_{\lambda',\lambda} \hat{W}^{\lambda}(x)$  *almost surely, with*  $c_{\lambda',\lambda} = e^{\Psi'(0+\int_{\lambda}^{\lambda'}$  $\lambda' \frac{du}{\Psi(u)}$ .

The process  $(\hat{W}^{\lambda}(x), x \ge 0)$  can be interpreted as follows. Define a random equivalence relation  $\mathscr A$  on  $(0, \infty)$  via

 $x \stackrel{\mathscr{A}}{\sim} y$  if and only if  $\hat{W}^{\lambda}(x) = \hat{W}^{\lambda}(y)$ .

This induces a random partition<sup>1</sup> of the set  $(0, \infty)$  into open intervals of constancy of  $\hat{W}^{\lambda}$ . A simple application of Lemma 3.6 ensures that this partition does not depend on  $\lambda$ . By definition, the subintervals of the partition  $\mathscr A$  are made of individuals whose ancestral lineages have the same asymptotic behaviour. These subintervals correspond to the jump intervals of  $W^{\lambda}$ , the subordinator obtained as the right-inverse of  $\hat{W}^{\lambda}$ , that is to say

$$
\mathscr{A} = \left\{ \left( W^{\lambda}(u-), W^{\lambda}(u) \right); u > 0 \right\} \text{ a.s.}
$$

In other words the families in  $\mathscr A$  are separated by points  $x_i, i \in I$ , in the support  $\mathscr S$  of the associated Stieltjes measure  $d\hat{W}^{\lambda}$ ,  $\mathscr{S} := \overline{\{W^{\lambda}(u) : u \geq 0\}}$ .

The next theorem states that  $\mathscr A$  corresponds actually to the families of current individuals having a common ancestor.

**Theorem 3.7.** *For any x*,  $y \in (0, \infty)$ ,

$$
x \stackrel{\mathscr{A}}{\sim} y
$$
 if and only if  $\hat{X}_t(x) = \hat{X}_t(y)$  for some  $t \ge 0$ .

Theorems 3.1 and 3.7 complete the results obtained under Grey's condition, in Foucart et al. [14, Sections 3, 4 and 5.3], on the long-term behaviour of the ancestral lineages, as well as on the representation of the ancestral partition when Grey's condition is not in force.

As stated in Theorem 3.1, there are two separate cases to treat according whether Grey's condition holds or not.

When  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} < \infty$  (Grey's condition), the process  $(\hat{W}^{\lambda}(x), x \ge 0)$  is the inverse of a compound Poisson process for any  $\lambda \in (0, \infty]$ . By taking  $\lambda = \infty$ , the latter has for jump measure the probability law  $v_{\infty}$  and the partition  $\mathscr A$  is thus constituted of i.i.d. families with lengths of law  $v_{\infty}$ , i.e.  $\mathscr A$  takes the form of a consecutive partition into intervals:

$$
\mathscr{A} = ((0, x_1), (x_1, x_2), \ldots)
$$
 a.s.,

where  $(x_i, i \ge 1)$  is a random renewal process with jump law  $v_{\infty}$ . An important example is the subcritical Feller diffusion, for which  $\Psi(q) = \sigma^2/2q^2 + \gamma q$  and  $\nu_{\infty}$  is the exponential law with parameter  $2\gamma/\sigma$ , see [22, Section 5.2]. According to [14, Theorem 4.2], the flow of ancestral lineage processes  $(\hat{X}_t, t \ge 0)$  has in this case the same law as a flow of supercritical Feller diffusions with constant immigration (at rate  $\sigma^2$ ). The almost-sure convergence stated in Theorem 3.1 can also be shown by applying Theorem 1 in Foucart et al. [15], where an almost-sure renormalisation of supercritical branching processes with immigration is provided. The following figure provides a schematic representation of the families, their lineages and

 $1$  Up to a negligible set.



Fig. 2. Schematic representation of ancestral families under Grey's condition.

the process  $\hat{W}^{\infty}$ , under Grey's condition. We mention that in the Feller case, only binary coalescences are possible. When  $\Psi$  is not the branching mechanism of the subcritical Feller diffusion, i.e.  $\pi \neq 0$ , coalescences between the ancestral lineages inside each family are possibly multiple and can be described using the notion of consecutive coalescents, see [14, Section 5].

When  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} = \infty$ , the description is more involved since the process  $(\hat{W}^{\lambda}(x), x \ge 0)$ has singular continuous paths and any fixed subinterval of  $(0, \infty)$  of finite length contains infinitely many small families with positive probability. Recall that  $\mathscr S$  is the support of the random singular measure  $d\hat{W}^{\lambda}$ .

**Proposition 3.8.**  $(\infty) := \lim_{u \to \infty} \frac{\Psi(u)}{u}$  ∈  $(0, \infty]$ *. For any x* > 0*, the Hausdorff dimension of*  $\mathscr{S} \cap [0, x]$  *is* 

$$
\dim_H(\mathcal{S} \cap [0, x]) = \frac{\Psi'(0+)}{\Psi'(\infty)} \in [0, 1) \ a.s.
$$
\n(3.15)

From (2.5), one sees that  $\Psi'(\infty) = \sigma^2 \cdot \infty + \gamma + \int_0^\infty x \pi(dx) \in (0, \infty]$ . In the case of a CSBP with unbounded variation, namely with  $\sigma > 0$  or  $\int_0^1 x\pi(dx) = \infty$ , one has  $\Psi'(\infty) = \infty$  and the Hausdorff dimension of  $\mathscr S$  is zero. In the bounded variation case, (3.15) can be rewritten as

$$
\dim_H(\mathscr{S} \cap [0, x]) = \frac{\gamma}{\gamma + \int_0^\infty h\pi(dh)} \text{ a.s.}
$$

**Remark 3.9.** If one denotes by  $PPP_{\lambda} := \sum_{i \in I} \delta_{(a_i^{\lambda}, \Delta_i^{\lambda})}$  the Poisson point process associated to the subordinator *W*<sup>λ</sup>, the atoms ( $\Delta_i^{\lambda}$ ,  $i \in I$ ) are by definition the sizes of the different families in  $\mathscr A$  and the atoms of jump times  $(a_i^{\lambda}, i \in I)$  are related to the rate of escape of the ancestral lineages and can be thought as some ancestral types. Let  $\epsilon > 0$ . The restriction of PPP<sub> $\lambda$ </sub> to the set  $(0, \infty) \times (\epsilon, \infty)$  takes the form  $PPP_{\lambda|(0,\infty)\times(\epsilon,\infty)} = \sum_{j=1}^{\infty} \delta_{(\mathfrak{S}^{\epsilon}_{j},D^{\epsilon}_{j})}$ , where  $(\mathfrak{S}^{\epsilon}_{j}, j \geq 1)$ is a sequence of i.i.d. exponential random variables with parameter  $\bar{\nu}_{\lambda}(\epsilon) := \nu_{\lambda}((\epsilon, \infty))$  and  $(D_j^{\epsilon}, j \geq 0)$  are i.i.d. random variables with law  $\frac{v_{\lambda}(dx)}{\bar{v}_{\lambda}(\epsilon)} \mathbb{1}_{(\epsilon,\infty)}(x)$ . The exponential random variables allow one to distinguish families between each others. If  $e_i^{\epsilon}$  is larger than  $e_j^{\epsilon}$ , then the ancestral lineage of the *j*th family with size larger than  $\epsilon$ , diverges slower than that of the *i*th. Note that when Grey's condition holds, the characteristic measure of the Poisson point process PPP<sub> $\lambda$ </sub> is finite and one can take  $\epsilon = 0$ . When Grey's condition fails, one needs first to restrict ourselves to families of size greater than  $\epsilon > 0$  in order to be able to rank the atoms of times of  $PPP<sub>\lambda</sub>$ . In the schematic representation given in Fig. 2, the divergence of the ancestral lineage of the family  $(x_2, x_3)$  is faster than those of  $(x_1, x_2)$  and  $(0, x_1)$ .

**Example 3.10.** Let  $\Psi$  be the branching mechanism with drift  $\gamma = 1$  and Lévy measure  $\pi(\text{d}x) = x^{-\alpha-1}e^{-x}\text{d}x$  with  $\alpha \in (0, 2)$ . Then by a Tauberian theorem, see e.g. Feller's book [12, Chapter XIII.5],

$$
\frac{\Psi(x)}{x} \sim \int_0^\infty (1 - e^{-xh}) h \bar{\pi}(h) \mathrm{d}h \sim \int_0^\infty c_\alpha x^{\alpha - 1},
$$

where for all  $h > 0$ ,  $\bar{\pi}(h) := \pi((h, \infty))$  and  $c_{\alpha}$  is a strictly positive constant.

- (i) If  $\alpha \in (1, 2)$ , then  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} < \infty$ , hence  $\Psi'(\infty) = \infty$ , and  $\mathscr{S}$  is a discrete set, thus  $\dim_H(\mathscr{S} \cap [0, x]) = 0$  a.s. for all  $x > 0$ .
- (ii) If  $\alpha = 1$  (Neveu case), then  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} = \infty$  and  $\Psi'(\infty) = \infty$ , hence  $\mathscr S$  is not a discrete set, but dim<sub>*H*</sub>( $\mathcal{S} \cap [0, x]$ ) = 0 a.s. for all  $x > 0$ .
- (iii) If  $\alpha$  < 1,  $\Psi'(\infty)$  <  $\infty$ , hence  $\int_{-\infty}^{\infty} \frac{du}{\Psi(u)}$  =  $\infty$  and  $\mathscr S$  is not a discrete set, and  $\dim_H(\mathscr{S} \cap [0, x]) = \frac{1}{1+\Gamma(1-\alpha)}$  a.s. for all  $x > 0$ .

In general, inverse subordinators do not have the Markov property. The joint density of the finite-dimensional marginals of  $(\hat{W}^{\lambda}(x), x \ge 0)$  is thus rather involved to study. We refer to the works of Lageras [20] and Veillette and Taqqu [32] for information on inverse subordinators.

The following proposition is a side result on the one-dimensional laws of the limiting process  $(\hat{W}^{\lambda}(x), x \ge 0)$ . Recall that  $\nu_{\lambda}$  denotes the Lévy measure of the subordinator  $(W^{\lambda}(x), x \ge 0)$ and  $\bar{\nu}_{\lambda}$  is its tail: for any  $x \geq 0$ ,  $\bar{\nu}_{\lambda}(x) := \nu_{\lambda}((x, \infty))$ .

**Proposition 3.11.** *The law of*  $\hat{W}^{\lambda}(x)$  *admits the density*  $g_x^{\lambda}$  *defined on*  $(0, \infty)$  *by* 

$$
g_x^{\lambda}(u) := \int_0^x \bar{\nu}_{\lambda}(x - z) \mathbb{P}(W^{\lambda}(u) \in dz).
$$
 (3.16)

*When*  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} < \infty$ ,  $\hat{W}^{\infty}(x)$  *has density*  $g_x^{\infty}$ :

$$
g_x^{\infty}(u) := e^{-u} \sum_{n=0}^{\infty} \frac{u^n}{n!} \int_0^x \bar{\nu}_{\infty}(x - z) \nu_{\infty}^{*n}(dz).
$$
 (3.17)

#### 4. Proofs

The most demanding part in the proof of Theorem 3.1 lies in the almost-sure convergence. The convergence in law will be established from a direct argument involving the subordinators  $(X_{-t,0}(x), x \ge 0)$  rather than their inverse, we refer to the forthcoming Lemma 4.14. In order to show the almost-sure convergence, we first establish a general result of independent interest relating  $\theta$ -invariant functions of stochastically monotone processes, with  $\theta$ -invariant measures of their dual processes, see Section 4.1. We then apply this result in the setting of CSBPs. The asymptotics of a certain  $\theta$ -invariant function for the dual process  $\hat{X}$  is studied. It enables to find a new martingale associated to the ancestral lineage process. We shall see how the renormalization  $v_t(\lambda)$  appears in this martingale and deduce the almost sure convergence of the process  $(v_t(\lambda) \hat{X}_t(x), t \ge 0)$  started from a fixed value *x*, see Lemma 4.10. We study the associated limiting process in *x* and show that it satisfies the properties stated in Theorem 3.1, see Lemma 4.11.

Theorem 3.7 will be established with the help of some previous results obtained in [14]. Proposition 3.8 will be a consequence of Theorem 3.1. The proof of Theorem 3.1 is divided into several lemmas. The first are needed to show the almost sure convergence towards some positive random variable  $\hat{W}^{\lambda}(x)$ .

#### *4.1. Invariant functions of stochastically monotone Markov processes*

In this section, we consider a "general" standard Markov process  $X := (X_t, t > 0)$  with state space  $[0, \infty)$ , and denote by  $(X_t(y), t \ge 0)$  the process started from  $y \in [0, \infty)$ . Recall that the process *X* is said to be stochastically monotone if for any  $t > 0$  and  $x \in [0, \infty)$ , the map  $y \mapsto \mathbb{P}(X_t(y) > x)$  is non-decreasing. Siegmund [29] has established that if the process *X* is stochastically monotone, non-explosive or with boundary  $\infty$  absorbing, and that for any fixed *t* and *z*, the map  $y \mapsto \mathbb{P}(X_t(y) \ge z)$  is right-continuous then there exists a unique Markov process  $\hat{X}$ , the so-called Siegmund dual process, such that for any *t* and *x*, *y* 

$$
\mathbb{P}(X_t(y) \ge x) = \mathbb{P}(\hat{X}_t(x) \le y). \tag{4.18}
$$

The latter identity can be rewritten as

$$
\mathbb{P}(\hat{X}_t(x) > y) = \mathbb{P}(x > X_t(y)).\tag{4.19}
$$

Our first result shows how to find fundamental martingales for the Siegmund dual process  $(\hat{X}_t(x), t \ge 0)$  of any stochastically monotone Markov process  $(X_t(x), t \ge 0)$ . Recall  $\hat{T}_y$  defined in (2.11).

**Theorem 4.1** (*Invariant Functions of*  $\hat{X}$ ). Let  $(P_t, t \ge 0)$  be the semigroup of the process  $(X_t, t \geq 0)$ *. Let*  $\theta \in \mathbb{R}$ *. If*  $\mu_{\theta}$  *is a positive Borel measure on*  $(0, \infty)$  *satisfying for any*  $t \geq 0$ ,  $\mu_{\theta}P_t = e^{\theta t}\mu_{\theta}$ , then the functions  $x \mapsto \mu_{\theta}([0, x))$  and  $x \mapsto \mu_{\theta}((x, \infty))$ , provided *they are well-defined, are*  $\theta$ *-invariant functions, namely functions*  $f_{\theta}$  *such that for any t*  $\geq 0$ *and*  $x \in [0, \infty)$ ,

$$
\mathbb{E}[f_{\theta}(\hat{X}_t(x))] = e^{\theta t} f_{\theta}(x),
$$

*so that*

$$
(e^{-\theta t} f_{\theta}(\hat{X}_t(x)), t \ge 0) \text{ is a martingale.}
$$
\n(4.20)

*In particular, if the process*  $(\hat{X}_t, t \ge 0)$  *has no positive jumps and*  $\mu_\theta$  *is finite on*  $[0, x)$  *for all*  $x > 0$ , then  $f_{\theta} : x \mapsto \mu_{\theta}([0, x))$  *is a well-defined increasing and left-continuous function, and for all*  $y \ge x \ge 0$ *,* 

$$
\mathbb{E}_x[e^{-\theta \hat{T}_y}] = \frac{\mu_\theta([0, x))}{\mu_\theta([0, y))}.
$$
\n(4.21)

**Remark 4.2.** A measure  $\mu_{\theta}$  satisfying  $\mu_{\theta} P_t = e^{\theta t} \mu_{\theta}$  is sometimes called an *eigen-measure* or a  $\theta$ -invariant measure.

**Remark 4.3.** Let  $\mathcal{L}$  be the generator of the process  $(X_t, t \ge 0)$ . The Kolmogorov forward equation entails that the condition  $\mu_{\theta} P_t = e^{\theta t} \mu_{\theta}$  for all  $t \ge 0$ , is equivalent to  $\mu_{\theta} \mathcal{L} = \theta \mu_{\theta}$ , where  $\mu_{\theta} \mathcal{L}$  is by definition the measure such that  $\langle \mu_{\theta} \mathcal{L}, f \rangle := \int \mathcal{L} f(x) \mu_{\theta}(\mathrm{d}x)$  for any function  $f \in C_b^2((0, \infty)).$ 

**Proof.** Set  $f_{\theta}(x) = \mu_{\theta}([0, x))$  for all  $x > 0$ . For any  $x \in (0, \infty)$  and any  $t \ge 0$ ,

$$
\hat{P}_t f_\theta(x) = \mathbb{E}\left[f_\theta\big(\hat{X}_t(x)\big)\right] = \mathbb{E}\left[\int \mathbb{1}_{\{\hat{X}_t(x) > y\}} \mu_\theta(\mathrm{d}y)\right]
$$
\n
$$
= \mathbb{E}\left[\int \mathbb{1}_{\{x > X_t(y)\}} \mu_\theta(\mathrm{d}y)\right] \text{ by the duality relation (4.19)}
$$
\n
$$
= \mu_\theta P_t\big([0, x)\big) = e^{\theta t} \mu_\theta\big([0, x)\big) = e^{\theta t} f_\theta(x).
$$

The martingale property (4.20) follows readily from the Markov property. Note that the map  $f_{\theta}: x \mapsto \mu_{\theta}([0, x))$  is left-continuous. We now apply the bounded optional stopping time theorem at time  $t \wedge \hat{T}_y$ :

$$
\mathbb{E}\big[e^{-\theta t\wedge \hat{T}_y}f_\theta\big(\hat{X}_{t\wedge \hat{T}_y}(x)\big)\big]=f_\theta(x).
$$

Since  $f_{\theta}$  is non-decreasing and  $\hat{X}_{t \wedge \hat{T}_y}(x) \leq y$  a.s, one has for any  $t \geq 0$ ,  $f_{\theta}(\hat{X}_{t \wedge \hat{T}_y}(x)) \leq f_{\theta}(y)$ . On the event  $\{\hat{T}_y < \infty\}$ , the left-continuity of  $f_\theta$  and the absence of negative jumps in the process  $(\hat{X}_t, t \ge 0)$  ensure that  $f_\theta\big(\hat{X}_{t \wedge \hat{T}_y}(x)\big) \longrightarrow_{t \to \infty} f_\theta(y)$ . This yields

$$
f_{\theta}(x) = \lim_{t \to \infty} \mathbb{E}\big[e^{-\theta t \wedge \hat{T}_y} f_{\theta}\big(\hat{X}_{t \wedge \hat{T}_y}(x)\big)\big] = \mathbb{E}\big[e^{-\theta \hat{T}_y} f_{\theta}(y) \mathbb{1}_{\{\hat{T}_y < \infty\}}\big],
$$

which provides the identity  $(4.21)$ .  $\Box$ 

Remark 4.4. Theorem 4.1 holds in general for any stochastically monotone Markov process. In particular, the process is not required to have one-sided jumps. The state space  $[0, \infty]$  could also be replaced by a more general nice ordered state space.

**Remark 4.5.** Let  $\hat{\mathcal{L}}$  denote the generator of the process  $\hat{X}$ . An invariant function  $f_{\theta}$  for the semigroup of  $\hat{X}$  can be thought as a solution to the equation  $\hat{\mathcal{L}} f_{\theta} = \theta f_{\theta}$ . However, when the process has jumps,  $\hat{\mathcal{L}}$  is an integro-differential operator and no general theory allows one for identifying solutions of this equation. Theorem 4.1 reveals that for stochastically monotone processes finding a  $\theta$ -invariant function corresponds to finding a  $\theta$ -invariant measure for the dual process. This is reminiscent to a result of  $Cox$  and Rösler [31].

#### *4.2. Application to CSBPs and proof of Theorem* 3.1

Recall the definition of  $(\hat{X}_t(x), t \ge 0)$  as the right-continuous inverse of  $(X_{-t,0}(x), t \ge 0)$ , Lemma 2.4 and the duality relation  $(2.13)$ . This relationship matches with  $(4.19)$  and we will be able to apply Theorem 4.1.

Recall the action (2.4) of the generator  $\mathcal L$  on exponential functions. We now look for the θ-invariant measures  $\mu_\theta$  in our setting and their Laplace transforms explicitly in terms of  $\Psi$ . The following Lemma holds for general branching mechanism (i.e. not necessarily subcritical).

**Lemma 4.6.** *For any*  $\theta > 0$ , the map  $c_{\theta} : q \mapsto e^{-\theta \int_1^q \frac{du}{\Psi(u)}}$  is the Laplace transform of a Borel *measure*  $\mu_{\theta}$  *on* [0,  $\infty$ )*. Moreover, the measure*  $\mu_{\theta}$  *is*  $\theta$ *-invariant for the semigroup* ( $P_t$ ,  $t \ge 0$ ) *of the CSBP*  $(X_t, t \geq 0)$ *.* 

**Proof.** Recall that L denotes the generator of the CSBP  $(X_t(x), t \ge 0)$ . Let  $\theta \ge 0$ . It is easily checked from the expression of  $c_\theta$  that  $(-1)^n c_\theta^{(n)} \ge 0$  on  $(0, \infty)$ . Bernstein theorem,

see e.g. [33, Theorem 12.b, page 161], guarantees that there exists a certain Borel measure  $\mu_{\theta}$ (possibly infinite) on  $[0, \infty)$  such that  $c_{\theta}(q) = \int_0^{\infty} e^{-qx} \mu_{\theta}(dx)$  for all  $q > 0$ . We now check that the measure  $\mu_{\theta}$  solves  $\mu_{\theta} \mathcal{L} = \theta \mu_{\theta}$ . Let  $q > 0$ , recall that  $\lambda > 0$  is fixed. Observe first that  $c_{\theta}$  satisfies the equation

$$
- \Psi(q)c_{\theta}'(q) = \theta c_{\theta}(q), \quad c_{\theta}(1) = 1.
$$
\n
$$
(4.22)
$$

Recall  $e_q : x \mapsto e^{-qx}$  for any  $x, q \ge 0$ . Since the linear span of exponential functions is a core for the generator  $\mathcal{L}$ , it is enough to verify that

$$
\langle \mu_{\theta} \mathcal{L}, e_{q} \rangle := \int_{0}^{\infty} \mathcal{L}e_{q}(x) \mu_{\theta}(\mathrm{d}x) = \langle \theta \mu_{\theta}, e_{q} \rangle = \theta \int_{0}^{\infty} e^{-qx} \mu_{\theta}(\mathrm{d}x). \tag{4.23}
$$

One has on the other hand  $\mathcal{L}e_q(x) = \Psi(q)x e_q(x)$  and (4.23) is equivalent to

$$
\Psi(q) \int_0^\infty x e^{-qx} \mu_\theta(\mathrm{d}x) = \theta \int_0^\infty e^{-qx} \mu_\theta(\mathrm{d}x),
$$
\nholds true by using (4.22).

which holds true by using  $(4.22)$ .  $\Box$ 

**Remark 4.7.** In the subcritical or critical cases, the map  $q \mapsto c_{\theta}(q)$  is completely monotone on (0,  $\infty$ ) and not defined at 0. This entails that the measure  $\mu_{\theta}$  is infinite. In the supercritical case,  $c_{\theta}$  is completely monotone *and* well-defined and right-continuous at 0, in this case the measure  $\mu_{\theta}$  is finite. Lastly, observe that  $\mu_{\theta}$  gives a mass at 0 if and only if Grey's condition holds (which ensures that  $c_{\theta}(\infty) > 0$ ).

According to Theorem 4.1, the map  $f_{\theta}$  :  $x \mapsto \mu_{\theta}([0, x))$  is a  $\theta$ -invariant function for  $(\hat{X}_t, t \ge 0)$ . The following simple calculation provides an expression of the Laplace transform of  $f_\theta$ . For any  $q > 0$ ,

$$
\xi_{\theta}(q) := \int_0^{\infty} f_{\theta}(y) e^{-qy} dy = \int_0^{\infty} \int_0^{\infty} \mathbb{1}_{\{u < y\}} e^{-qy} \mu_{\theta}(du) dy = \frac{1}{q} c_{\theta}(q). \tag{4.25}
$$

Inverting  $\xi_{\theta}$  in order to find  $f_{\theta}$  does not seem to be feasible in a general setting, however we shall see in the next lemma that  $\xi_{\theta}$  has regular variation properties at 0, Tauberian theorems will then allow us to find an equivalent at  $\infty$  of the function  $f_\theta$  and hence enable us to investigate more precisely the martingale  $(e^{-\theta t} f_{\theta}(\hat{X}_t(x)), t \ge 0)$ . Arguments for establishing the three next lemmas are lifted from those in Pakes [28, Theorem 10] and Barbour [2] and adapted to the setting of continuous-state space branching processes.

**Lemma 4.8.** Assume  $\Psi'(0+) \neq 0$ . The map  $R: q \mapsto e^{-\int_{1}^{q} \frac{du}{\Psi(u)}}$  is regularly varying at 0 *with index*  $-1/\Psi'(0+)$ *. In particular it takes the form*  $R(q) = q^{-\frac{1}{\Psi'(0+)}} L_1(1/q)$ *<i>, where*  $L_1$  *is a slowly varying function at*  $\infty$ *. Moreover, for any*  $\theta > -\Psi'(0+)$ *,* 

$$
f_{\theta}(y) \underset{y \to \infty}{\sim} y^{\frac{\theta}{\Psi'(0+)}} \frac{1}{\Gamma\left(1 + \frac{\theta}{\Psi'(0+)}\right)} L_1(y)^{\theta} = \frac{1}{\Gamma\left(1 + \frac{\theta}{\Psi'(0+)}\right)} R(1/y)^{\theta}.
$$
 (4.26)

**Proof.** For any  $q > 0$ ,

$$
\int_{q}^{1} \frac{du}{\Psi(u)} = \int_{q}^{1} \left( \frac{1}{\Psi(u)} - \frac{1}{\Psi'(0+)u} \right) du + \int_{q}^{1} \frac{du}{\Psi'(0+)u}
$$

$$
= \int_{1}^{1/q} \left( \frac{1}{u^2 \Psi(1/u)} - \frac{1}{\Psi'(0+)u} \right) du - \frac{1}{\Psi'(0+)} \log(q).
$$

Set  $\epsilon(u) = \frac{1}{u \Psi(1/u)} - \frac{1}{\Psi'(0+)}$  for  $u > 0$  and notice that  $\epsilon(u) \longrightarrow 0$ . One has

$$
R(q) = q^{-\frac{1}{\Psi'(0+)}} \exp\left(\int_1^{1/q} \frac{\epsilon(u)}{u} du\right)
$$

and by Karamata's representation theorem, see e.g. Bingham et al. [7, Theorem 1.3.1], we see that  $L_1(x) := \exp\left(\int_1^x \frac{\epsilon(u)}{u} du\right)$  is slowly varying at  $\infty$ . Recall  $c_\theta(q) = e^{-\theta \int_1^q \frac{du}{\Psi(u)}}$ . By (4.25),  $\xi_{\theta}: q \mapsto \frac{c_{\theta}(q)}{q}$  $\frac{q}{q}$  is regularly varying at 0 with index  $\rho = -1 - \frac{\theta}{\Psi'(0+)}$ . The Tauberian theorem with monotone density, see e.g. [12, Chapter XIII.5, Theorem 4], provides (4.26).

From now on, we focus on the subcritical case  $\Psi'(0+) > 0$ . We now establish an almost sure limit theorem for the first passage time above level *y*,  $\hat{T}_y^x$ , when *y* goes to  $\infty$ .

**Lemma 4.9.** *For any*  $x > 0$ *, almost-surely* 

$$
\hat{T}_y^x - \int_{1/y}^1 \frac{\mathrm{d}u}{\Psi(u)} \underset{y \to \infty}{\longrightarrow} S(x),
$$

*where S*(*x*) *is a finite random variable with Laplace transform*

$$
\mathbb{E}[e^{-\theta S(x)}] = \Gamma\left(1 + \frac{\theta}{\Psi'(0+)}\right) f_{\theta}(x) \text{ for all } \theta > -\Psi'(0+). \tag{4.27}
$$

**Proof.** We establish first a convergence in law. By the asymptotic equivalence  $(4.26)$ , for any  $\theta > -\Psi'(0+),$ 

$$
\mathbb{E}[e^{-\theta(\hat{T}_y^x-\int_{1/y}^1 \frac{du}{\psi(u)})}] \longrightarrow_{y \to \infty} \Gamma\left(1+\frac{\theta}{\psi'(0+)}\right) f_{\theta}(x).
$$

One obtains by applying a continuity theorem for the moment generating function, see [10, Theorem 3], that as *y* goes to  $\infty$ ,

$$
\hat{T}_y^x - \int_{1/y}^1 \frac{\mathrm{d}u}{\Psi(u)} \xrightarrow{\mathcal{L}} S(x),\tag{4.28}
$$

where  $S(x)$  has Laplace transform  $(4.27)$ . We now show the almost sure convergence. Let  $(y_n, n \ge 0)$  be an increasing sequence such that  $y_n \rightarrow \infty$  and with  $y_0 := x$ . Write

$$
\hat{T}_{y_n}^x - \int_{1/y_n}^{1/x} \frac{du}{\Psi(u)} = \sum_{k=0}^n \left( \hat{T}_{y_{k+1}}^x - \hat{T}_{y_k}^x - \int_{1/y_{k+1}}^{1/y_k} \frac{du}{\Psi(u)} \right).
$$

Since the Markov process  $\hat{X}$  has no positive jumps, the summands are independent and by (4.28), the series on the right-hand side converges in law, one can apply [8, Theorem 9.5.5], which ensures that the series actually converges almost-surely. We finally get that

$$
\hat{T}_y^x - \int_{1/y}^1 \frac{\mathrm{d}u}{\Psi(u)} \underset{y \to \infty}{\longrightarrow} S(x) \text{ a.s. } \square
$$

We now deduce the convergence of the process  $(v_t(\lambda)\hat{X}_t(x), t \ge 0)$  for fixed *x* and a representation of its limit  $\hat{W}^{\lambda}(x)$ .

**Lemma 4.10.** Let  $\lambda > 0$ . For any  $x > 0$ , the following almost sure convergence holds:

$$
v_t(\lambda)\hat{X}_t(x) \underset{t \to \infty}{\longrightarrow} \hat{W}^{\lambda}(x) := c(\lambda)e^{-\Psi'(0+)S(x)},
$$
\n(4.29)

*where*  $c(\lambda)$  *is a constant independent of x.* 

**Proof.** We first establish the almost sure convergence towards some random variable  $\hat{W}^{\lambda}(x)$ . By applying Theorem 4.1, the process  $(e^{-t} f_1(\hat{X}_t(x)), t \ge 0)$  is a positive martingale. Therefore, the latter converges almost surely towards some random variable  $Z(x)$ . Recall Lemma 4.8. Set  $\beta(1,\lambda) := \frac{\lambda \xi_1(\lambda)}{\Gamma(1+1/\Psi'(0+))}$  and  $R_{\lambda}(1/y) := \exp\left(\int_{1/y}^{\lambda} \frac{du}{\Psi(u)}\right) = \exp\left(\int_{1}^{\lambda} \frac{du}{\Psi(u)}\right) R(1/y)$ . Recall that  $\hat{X}_t(x) \longrightarrow_{t \to \infty} \infty$  a.s, see Proposition 2.3. By Lemma 4.8,

$$
e^{-t} f_1(\hat{X}_t(x)) \underset{t\to\infty}{\sim} e^{-t} \beta(1,\lambda) R_{\lambda}(1/\hat{X}_t(x)) \underset{t\to\infty}{\longrightarrow} Z(x)
$$
 a.s.

Hence

$$
R_{\lambda}\left(1/\hat{X}_t(x)\right) \underset{t\to\infty}{\sim} \frac{e^t Z(x)}{\beta(1,\lambda)}.
$$
\n(4.30)

Now, using the fact that  $R_{\lambda}$  is non-decreasing and regularly varying at 0 with index  $-\frac{1}{\psi'(0+)}$ , we see that  $R_{\lambda}^{-1}$  is regularly varying at  $\infty$  with index  $-\Psi'(0+)$ . Taking  $R_{\lambda}^{-1}$  in the asymptotic equivalence (4.30) yields

$$
\frac{1}{\hat{X}_t(x)} \underset{t \to \infty}{\sim} R_{\lambda}^{-1} \left( \frac{e^t Z(x)}{\beta(1,\lambda)} \right) \underset{t \to \infty}{\sim} \left( \frac{Z(x)}{\beta(1,\lambda)} \right)^{-\Psi'(0+)} R_{\lambda}^{-1}(e^t).
$$

By the definition of  $\lambda \mapsto v_t(\lambda)$ , see (2.7), one has  $R_\lambda(v_t(\lambda)) = e^t$ . This allows us to conclude that

$$
v_t(\lambda)\hat{X}_t(x) \xrightarrow[t \to \infty]{} \hat{W}^{\lambda}(x) := \left(Z(x)/\beta(1,\lambda)\right)^{\Psi'(0+)} \text{ a.s.}
$$

We now explain the relation between  $\hat{W}^{\lambda}(x)$  and the random variable  $S(x)$  introduced in the previous lemma. Denote by  $(F_t)_{t\geq 0}$  the natural filtration of  $\hat{X}$ . Let  $(M_t, t \geq 0)$  be the martingale defined by  $M_t := \mathbb{E}[e^{-S(x)}|\mathcal{F}_t]$  for any  $t \ge 0$ . Note that  $M_t \longrightarrow e^{-S(x)}$  almost surely. Moreover by Lemma 4.9, and Lebesgue's theorem, for any  $t \ge 0$ ,

$$
M_t = \lim_{y \to \infty} \mathbb{E}\left[e^{-\left(\hat{T}_y^x - \int_{1/y}^1 \frac{du}{\Psi(u)}\right)} | \mathcal{F}_t\right].
$$

For any fixed y, conditional on  $\{\hat{T}_y^x > t\}$ , we have that  $\hat{T}_y^x = \hat{T}_y^x \circ \theta_t + t$  with  $\hat{T}_y^x \circ \theta_t := \inf\{s > t\}$  $0: \hat{X}_{s+t}(x) > y$ . By the Markov property, we get almost surely

$$
\mathbb{E}[e^{-\hat{T}_y^x}|\mathcal{F}_t] = \frac{f_1\left(\hat{X}_{t\wedge \hat{T}_y^x}(x)\right)}{f_1(y)}e^{-t\wedge \hat{T}_y^x}.
$$

Therefore, almost surely

$$
M_{t} = \lim_{y \to \infty} \frac{e^{\int_{1/y}^{1} \frac{du}{\Psi(u)}}}{f_{1}(y)} f_{1} \left(\hat{X}_{t \wedge \hat{T}_{y}^{x}}(x)\right) e^{-t \wedge \hat{T}_{y}^{x}}
$$
  
=  $\Gamma \left(1 + \frac{1}{\Psi'(0+)}\right) f_{1}(\hat{X}_{t}(x)) e^{-t}.$ 

Hence,

$$
Z(x) := \lim_{t \to \infty} e^{-t} f_1(\hat{X}_t(x)) = \frac{1}{\Gamma\left(1 + \frac{1}{\Psi'(0+)}\right)} e^{-S(x)} > 0 \text{ a.s.},
$$

and we have  $\hat{W}^{\lambda}(x) = c(\lambda)e^{-\Psi'(0+)S(x)}$  a.s.  $\square$ 

The almost sure convergence in Lemma  $4.10$  holds for fixed  $x$ . We now further investigate the limiting process in the variable  $x$  and establish a stronger convergence.

**Lemma 4.11.** The process  $(\hat{W}^{\lambda}(x), x \ge 0)$  admits a non-decreasing right-continuous *modification. Moreover, setting*  $J^{\lambda} := \{x > 0 : \hat{W}^{\lambda}(x) > \hat{W}^{\lambda}(x-\})\}$ , one has almost surely,

$$
\forall x \notin J^{\lambda}, \ v_t(\lambda)\hat{X}_t(x) \underset{t \to \infty}{\longrightarrow} \hat{W}^{\lambda}(x). \tag{4.31}
$$

**Proof.** For technical reasons it will be easier to establish the result first with a left-continuous limiting random process instead of the targeted right-continuous one. Once a version with left-continuous paths and right limits is constructed, one can easily consider the associated right-continuous version of it. Recall that by Lemma 4.10, for any fixed *x*,  $v_t(\lambda)\hat{X}_t(x)$  converges almost surely as *t* goes to  $\infty$  towards a random variable denoted by  $\hat{W}^{\lambda}(x)$ . Consider the almost sure event  $\Omega_1$  on which all random variables  $(\hat{W}^{\lambda}(q), q \in \mathbb{Q}_+)$  are defined. Since for any  $t \ge 0$ , the process  $(\hat{X}_t(x), x \ge 0)$  is non-decreasing then for any rational numbers  $q' \ge q \ge 0$ , one has  $\hat{W}^{\lambda}(q') \ge \hat{W}^{\lambda}(q)$ . We work deterministically on  $\Omega_1$  and define for all  $x \ge 0$ ,

$$
\tilde{\hat{W}}^{\lambda}(x) := \lim_{\substack{q \uparrow x, \\ q \in \mathbb{Q}_+}} \hat{W}^{\lambda}(q).
$$

The process  $(\tilde{W}^{\lambda}(x), x \ge 0)$  is well-defined on  $\Omega_1$  and we define it as the null process on  $\Omega \setminus \Omega_1$ . By construction,  $(\hat{\hat{W}}^{\lambda}(x), x \ge 0)$  is left-continuous. It has right limits since it is non-decreasing. We now show that for any  $x \ge 0$ ,  $\mathbb{P}(\hat{W}^{\lambda}(x)) = \tilde{W}^{\lambda}(x) = 1$ . By the identity (4.29), on the event  $\Omega_1$ , we have that for all  $q \in \mathbb{Q}_+$  such that  $q < x$ 

$$
\hat{W}^{\lambda}(q) = \hat{W}^{\lambda}(x)e^{-\Psi'(0+)(S(q)-S(x))} \text{ a.s.}
$$
\n(4.32)

We now verify that  $S(q) \longrightarrow q \to x \atop q < x$ *S*(*x*) a.s. Since almost surely for all  $x \le q \le q^{\prime}$ ,  $\hat{X}_t(x) \le$  $\hat{X}_t(q) \leq \hat{X}_t(q')$ , then  $\hat{T}_y^{q'} \leq \hat{T}_y^q \leq \hat{T}_y^x$  and similarly, almost surely  $S(q') \leq S(q) \leq S(x)$ .

Hence  $(S(q), q > 0)$  is decreasing. Recall the Laplace transform of  $S(q)$  given in (4.27), since the function  $f_{\theta}: x \mapsto \mu_{\theta}([0, x))$  is left-continuous, we have that

$$
\mathbb{E}[e^{-\theta S(q)}] \underset{q < x}{\longrightarrow} \mathbb{E}[e^{-\theta S(x)}].
$$

Finally we see that  $S(q) \longrightarrow q \to x \atop q < x$ *S*(*x*) a.s. and by the identity (4.32),  $\hat{W}^{\lambda}(q) \longrightarrow q^{\lambda} \longrightarrow q^{\lambda} \hat{W}^{\lambda}(x)$ , hence  $\hat{W}^{\lambda}(x) = \tilde{\hat{W}}^{\lambda}(x)$  a.s. and  $\tilde{\hat{W}}^{\lambda}$  is a left-continuous version of the family of random variables  $(\hat{W}^{\lambda}(x), x \ge 0)$ . If we now define on  $\Omega_1$ , simultaneously for all *x*, the process  $(\hat{W}^{\lambda}(x), x \ge 0)$ by setting  $\hat{W}^{\lambda}(x) := \hat{W}^{\lambda}(x+)$ , then the process  $(\hat{W}^{\lambda}(x), x \ge 0)$  is right-continuous. The

first statement of the lemma is established. It remains to see that the almost sure pointwise convergence holds outside the set of jumps  $J^{\lambda}$ . Almost surely for all  $x \notin J^{\lambda}$ , one can choose two rational numbers *q* and *q'* such that  $q' < x < q$ . Since  $\hat{X}_t(q') \leq \hat{X}_t(x) \leq \hat{X}_t(q)$  for all *t*, one has

$$
\widetilde{\hat{W}}^{\lambda}(q') \leq \liminf_{t \to \infty} v_t(\lambda) \hat{X}_t(x) \leq \limsup_{t \to \infty} v_t(\lambda) \hat{X}_t(x) \leq \widetilde{\hat{W}}^{\lambda}(q).
$$

Since  $\tilde{W}^{\lambda}$  has left-continuous paths with right limits and  $x \notin J^{\lambda}$ , both sides of the inequalities above converge towards the same value  $\tilde{\hat{W}}^{\lambda}(x)$  when  $q' \uparrow x$  and  $q \downarrow x$ . By definition of  $\hat{W}^{\lambda}(x)$ , since  $x \notin J^{\lambda}$ ,  $\hat{W}^{\lambda}(x) = \tilde{\hat{W}}^{\lambda}(x)$ . This allows us to claim (4.31).  $\square$ 

From now on we work with the right-continuous version of  $\hat{W}^{\lambda}$ . The next lemma sheds some light on the role of the parameter  $\lambda$  and provides Lemma 3.6.

**Lemma 4.12.** *For any*  $\lambda > 0$  *and*  $\lambda' > 0$ *,* 

$$
\frac{v_t(\lambda)}{v_t(\lambda')} \underset{t \to \infty}{\longrightarrow} c_{\lambda, \lambda'} := e^{\Psi'(0 + )\int_{\lambda'}^{\lambda} \frac{du}{\Psi(u)}}.
$$
\n(4.33)

*Moreover*  $\hat{W}^{\lambda}(x) = c_{\lambda,\lambda'} \hat{W}^{\lambda'}(x)$  *for all*  $x \in (0, \infty)$  *almost surely.* 

**Proof.** Since  $\Psi'(0+) \ge 0$ ,  $v_t(\lambda) \longrightarrow_{t \to \infty} 0$ . Moreover,  $\lim_{t \to \infty} \uparrow \frac{\Psi(v_t(u))}{v_t(u)} = \Psi'(0+)$ . Recall that  $\frac{d}{d\lambda}v_t(\lambda) = \frac{\Psi(v_t(\lambda))}{\Psi(\lambda)}$ . Therefore for any  $\lambda \neq \lambda'$ ,

$$
\frac{v_t(\lambda)}{v_t(\lambda')} = \exp\left(\int_{\lambda'}^{\lambda} \frac{d}{du} \log v_t(u) du\right) = \exp\left(\int_{\lambda'}^{\lambda} \frac{\Psi(v_t(u))}{v_t(u)} \frac{du}{\Psi(u)}\right)
$$

and by monotone convergence  $\frac{v_t(\lambda)}{v_t(\lambda')} \longrightarrow_{t \to \infty} \exp\left(\frac{\Psi'(0+)}{\lambda'} \int_{\lambda'}^{\lambda} \frac{du}{\Psi(u)}\right)$ . We see from Lemma 4.11 that  $\hat{W}^{\lambda}(x) = c_{\lambda,\lambda'} \hat{W}^{\lambda'}(x)$  for all  $x \in (0,\infty)$  almost surely.  $\square$ 

**Lemma 4.13.** *The map*  $\kappa_{\lambda}$  :  $q \mapsto e^{-\Psi'(0+) \int_q^{\lambda} \frac{du}{\Psi(u)}}$  is the Laplace exponent of a drift-free *subordinator*  $W^{\lambda}$ *. Its Lévy measure, denoted by*  $v_{\lambda}$ *, is finite if and only if*  $\int^{\infty} \frac{du}{\Psi(u)} < \infty$ *.* 

**Proof.** For any  $\theta > 0$  and any  $y \in (0, \infty)$ , since  $X_{-t,0}(z)$  has the same law as  $X_{0,t}(z)$  for all *z* and  $t \geq 0$ , one gets by (2.6) and by applying Lemma 4.12

$$
\mathbb{E}[e^{-\theta X_{-t,0}\left(\frac{y}{v_t(\lambda)}\right)}] = e^{-y\frac{v_t(\theta)}{v_t(\lambda)}} \underset{t \to \infty}{\longrightarrow} e^{-y\kappa_\lambda(\theta)}.
$$
\n(4.34)

At any time *t*, the process  $(X_{-t,0}(y/v_t(\lambda)), y \ge 0)$  is a subordinator, the function  $\kappa_{\lambda}$  is therefore the Laplace exponent of a certain subordinator ( $W^{\lambda}(y)$ ,  $y \ge 0$ ). We show that there is no drift in the subordinator. Recall  $\Psi'(\infty) := \lim_{u \to \infty} \frac{\Psi(u)}{u} \in (0, \infty]$ . Since the case of linear branching mechanism is discarded, i.e  $\Psi(q) \neq bq$ , one has by convexity,  $\Psi'(\infty) > \Psi'(0+)$ . Choose  $\delta \in (\Psi'(0+), \Psi'(\infty))$ . There exists  $\lambda_0$  such that for all  $u \ge \lambda_0$ ,  $\frac{\Psi(u)}{u} \ge \delta$  and thus  $\frac{1}{\Psi(u)} \le \frac{1}{\delta u}$ . Therefore

$$
\lim_{\theta \to \infty} \int_{\lambda_0}^{\theta} \left( \frac{1}{\Psi'(0+)u} - \frac{1}{\Psi(u)} \right) du \ge \int_{\lambda_0}^{\infty} \left( \frac{1}{\Psi'(0+)} - \frac{1}{\delta} \right) \frac{du}{u} = \infty.
$$

One deduces that

$$
\frac{\kappa_{\lambda}(\theta)}{\theta} = \frac{1}{\lambda} \exp\left(-\Psi'(0+) \int_{\lambda}^{\theta} \left(\frac{1}{\Psi'(0+)u} - \frac{1}{\Psi(u)}\right) du\right) \underset{\theta \to \infty}{\longrightarrow} 0,
$$

which entails that there is no drift. Letting  $\theta$  to  $\infty$  in  $\kappa_{\lambda}(\theta)$ , we see that

$$
\lim_{\theta \to \infty} \kappa_{\lambda}(\theta) = \nu_{\lambda}((0, \infty)) = e^{\Psi'(0+) \int_{\lambda}^{\infty} \frac{du}{\Psi(u)}},
$$

which is finite if and only if  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} < \infty$ . Therefore the Lévy measure  $v_{\lambda}$  is finite if and only if  $\int_{0}^{\infty} \frac{du}{\psi(u)} < \infty$ . □

**Lemma 4.14.** *The process*  $(\hat{W}^{\lambda}(x), x \ge 0)$  *has the same finite-dimensional law as*  $((W^{\lambda})^{-1}(x), x \ge 0),$ 

*khere W*<sup>λ</sup> is a càdlàg subordinator with Laplace exponent  $\kappa$ <sub>λ</sub> and

 $(W^{\lambda})^{-1}(x) := \inf\{y \ge 0 : W^{\lambda}(y) > x\}.$ 

*Moreover, if*  $\int_{-\infty}^{\infty} \frac{du}{\Psi(u)} = \infty$ , then the process  $\hat{W}^{\lambda}$  has continuous paths almost surely.

**Proof.** By independence and stationarity of the increments of  $(X_{-t,0}(y/v_t(\lambda)), y \ge 0)$ , the convergence in law of the one-dimensional marginal  $(X_{-t,0}(y/v_t(\lambda)), t \ge 0)$  as t goes to  $\infty$  towards  $W^{\lambda}(y)$ , established in (4.34), entails the convergence of the finite-dimensional marginals. Since there is no drift part in the Laplace exponent  $\kappa_{\lambda}$ , the range of the subordinator *W*<sup> $\lambda$ </sup> contains almost surely no fixed point, see [3, Proposition 1.9-(i)]. Hence for any *x* ∈ (0, ∞) and  $y \in (0, \infty)$ ,  $\mathbb{P}(W^{\lambda}(y) = x) = 0$ . The weak convergence of the finite-dimensional marginals of  $X_{-t,0}(y/v_t(\lambda), y \ge 0)$  entails thus that for any  $0 < y_1 < \cdots < y_n$  and  $0 < x_1 < \cdots < x_n$ 

$$
\lim_{t\to\infty} \mathbb{P}\big(X_{-t,0}(y_1/v_t(\lambda))\geq x_1, X_{-t,0}(y_2/v_t(\lambda))\geq x_2, \ldots, X_{-t,0}(y_n/v_t(\lambda))\geq x_n\big)
$$
\n
$$
= \mathbb{P}\big(W^{\lambda}(y_1)\geq x_1, W^{\lambda}(y_2)\geq x_2, \ldots, W^{\lambda}(y_n)\geq x_n\big).
$$

By applying Lemma 4.10, the definition of the flow  $(\hat{X}_t(x), x \ge 0)$  and the duality relation (2.12), one gets the following identities

$$
\mathbb{P}\left(\hat{W}^{\lambda}(x_{1}) \leq y_{1}, \ \hat{W}^{\lambda}(x_{2}) \leq y_{1}, \ \ldots, \ \hat{W}^{\lambda}(x_{n}) \leq y_{n}\right) \tag{4.35}
$$
\n
$$
= \lim_{t \to \infty} \mathbb{P}\left(v_{t}(\lambda)\hat{X}_{t}(x_{1}) \leq y_{1}, v_{t}(\lambda)\hat{X}_{t}(x_{2}) \leq y_{2}, \ \ldots, v_{t}(\lambda)\hat{X}_{t}(x_{n}) \leq y_{n}\right)
$$
\n
$$
= \lim_{t \to \infty} \mathbb{P}\left(X_{-t,0}(y_{1}/v_{t}(\lambda)) \geq x_{1}, \ X_{-t,0}(y_{2}/v_{t}(\lambda)) \geq x_{2}, \ \ldots, \ X_{-t,0}(y_{n}/v_{t}(\lambda)) \geq x_{n}\right)
$$
\n
$$
= \mathbb{P}\left(W^{\lambda}(y_{1}) \geq x_{1}, \ W^{\lambda}(y_{2}) \geq x_{2}, \ \ldots, \ W^{\lambda}(y_{n}) \geq x_{n}\right)
$$
\n
$$
= \mathbb{P}\left((W^{\lambda})^{-1}(x_{1}) \leq y_{1}, \ (W^{\lambda})^{-1}(x_{2}) \leq y_{2}, \ \ldots, \ (W^{\lambda})^{-1}(x_{n}) \leq y_{n}\right).
$$
\n(4.35)

The processes  $(W^{\lambda})^{-1}$  and  $\hat{W}^{\lambda}$  have therefore the same law. The fact that when  $\int^{\infty} \frac{du}{\Psi(u)} = \infty$ , there are no jumps, i.e  $J^{\lambda} = \emptyset$  a.s., comes from the fact that the process  $(W^{\lambda})^{-1}$  is the inverse of a subordinator with an infinite Lévy measure. Since there is no drift in  $W^{\lambda}$ , the sample paths of  $\hat{W}^{\lambda}$  are pure singular continuous functions.  $\square$ 

**Proof of Theorem 3.1.** The main theorem is obtained by combining Lemmas 4.11, 4.14 and 4.13. □

**Remark 4.15.** When  $\int_{-\infty}^{\infty} \frac{du}{\psi(u)} = \infty$ , for any time  $t > 0$ , the subordinator  $(X_{-t,0}(x), x \ge 0)$  has an infinite Lévy measure, the process  $(\hat{X}_t(x), x \ge 0)$  is therefore continuous increasing. Since its limit  $(\hat{W}^{\lambda}(x), x \ge 0)$  is continuous, Dini's theorems ensure that the almost sure convergence (4.31) holds true locally uniformly.

.

**Proof of Corollary 3.3.** Recall the statement of Corollary 3.3. For any  $\lambda > 0$ , by (2.7), one has for any time *t*,

$$
\int_{v_I(\lambda)}^{\lambda} \left( \frac{1}{\Psi'(0+)u} - \frac{1}{\Psi(u)} \right) du = \frac{1}{\Psi'(0+)} \log \left( \frac{\lambda}{v_I(\lambda)} \right) - t
$$

$$
= \frac{1}{\Psi'(0+)} \log \left( \frac{\lambda}{v_I(\lambda)} e^{-\Psi'(0+)t} \right)
$$

Recall that  $v_t(\lambda) \longrightarrow 0$ . Therefore, the asymptotics  $v_t(\lambda) \sim c_\lambda e^{-\Psi'(0+)t}$  holds for some constant  $c_{\lambda} > 0$  if and only if

$$
\int_0^\lambda \left( \frac{1}{\Psi'(0+)u} - \frac{1}{\Psi(u)} \right) du < \infty. \tag{4.36}
$$

One has, when (4.36) holds,  $c_{\lambda} = \lambda e^{-\frac{\Psi'(0+) \int_0^{\lambda} (\frac{1}{\Psi'(0+)u} - \frac{1}{\Psi(u)}) du}}$ . Since  $\Psi(u) \sim \Psi'(0+)u$ , the convergence (4.36) is equivalent to  $\int_0^{\lambda} \left( \frac{\Psi(u) - \gamma u}{u^2} \right)$  $\left(\frac{\mu}{u^2}\right)$  du <  $\infty$ , where we recall that  $\gamma = \Psi'(0)$ is the linear drift in  $\Psi$ , see (2.5). Simple calculations from the Lévy–Khintchine form (2.5)

ensure that the latter integral converges if and only if  $\int_0^\infty u \log u \pi(du) < \infty$ . We refer for instance to the calculations around Proposition 3.14 in [26]. By Theorem 3.1, almost surely for all  $x \notin J^{\lambda}$ ,

$$
e^{-\Psi'(0+)t}\hat{X}_t(x) \underset{t\to\infty}{\longrightarrow} \frac{1}{c_{\lambda}}\hat{W}^{\lambda}(x).
$$

The process  $(\frac{1}{c_{\lambda}}\hat{W}^{\lambda}(x), x \ge 0)$  is the inverse of the subordinator  $(W^{\lambda}(c_{\lambda}x), x \ge 0)$ , whose Laplace exponent is  $\theta \mapsto c_{\lambda} \kappa_{\lambda}(\theta)$ . Recall  $\kappa_{\lambda}$ , one easily checks that

$$
c_{\lambda} \kappa_{\lambda}(\theta) = \lambda e^{-\Psi'(0+)\left[\int_0^{\lambda} \left(\frac{1}{\Psi'(0+\mu)} - \frac{1}{\Psi(u)}\right) du - \int_{\lambda}^{\theta} \frac{du}{\Psi(u)}\right]} = \theta e^{-\Psi'(0+)\int_0^{\theta} \left(\frac{1}{\Psi'(0+\mu)} - \frac{1}{\Psi(u)}\right) du}.
$$

#### *4.3. Proof of Theorem* 3.7

We establish now that the random partition  $\mathscr A$  matches with the ancestral partition. The proof relies on discretizations of the current population along the sequence of jumps times of a Poisson process, see [14, Section 5.2]. Consider a Poisson point process  $P$  on  $\mathbb{R}_+ \times \mathbb{R}_+$  with intensity the Lebesgue measure on  $\mathbb{R}_+ \times \mathbb{R}_+$ . Let  $\mu > 0$  and  $(J_j^{\mu}, j \ge 1)$  be the sequence of atoms (i.e. of jump times) of the homogeneous Poisson process ( $\mathcal{P}([0, \mu] \times [0, x])$ ,  $x \ge 0$ ). Note that all jump times of  $(\mathcal{P}([0, \mu] \times [0, x]), x \ge 0)$  are jump times of  $(\mathcal{P}([0, \mu'] \times [0, x]), x \ge 0)$ when  $\mu' > \mu$ . Therefore the sequence  $(J_i^{\mu'}$  $j^{\mu'}$ ,  $j \ge 1$ ) is thinner than  $(J_j^{\mu}, j \ge 1)$ . Moreover the set of all jump times of P,  $\mathcal{M} := \{J_i^{\hat{\mu}}, i \geq 1, \mu > 0\}$ , is almost surely an everywhere dense subset of  $[0, \infty)$ . One samples now individuals in the current population along the sequence  $(J_i^{\mu}, i \ge 1)$  for a fixed  $\mu$ . Let  $i \ne j \in \mathbb{N}$ . We first establish that the individuals  $J_i^{\mu}$  and  $J_j^{\mu}$  have the same ancestors if and only if  $J_i^{\mu}$  $\int_{0}^{\infty} J_j^{\mu}$  i.e.  $\hat{W}^{\lambda}(J_i^{\mu}) = \hat{W}^{\lambda}(J_j^{\mu})$ . The first implication is obvious since by definition of  $\hat{W}^{\lambda}$ , if  $\hat{X}_t(J_i^{\mu}) = \hat{X}_t(J_j^{\mu})$  for some  $t \ge 0$ , then  $\hat{W}^{\lambda}(J_i^{\mu}) = \hat{W}^{\lambda}(J_j^{\mu})$ . Denote by  $(C^{\mu}(t), t \ge 0)$  the process defined as follows

$$
i \stackrel{C^{\mu}(t)}{\sim} j
$$
 if and only if  $\hat{X}_t(J_i^{\mu}) = \hat{X}_t(J_j^{\mu})$  for any  $i, j \in \mathbb{N}^*$ .

According to [14, Proposition 4.18] and its proof, the process ( $C^{\mu}(t)$ ,  $t \ge 0$ ), called consecutive coalescent in [14], admits an almost-sure limit  $C^{\mu}(\infty)$  and for any fixed integers *i* and *j*,

*i*  $\int_0^{t} \int_0^{t}$  if and only if there exists a time *t*<sub>*i*,*j*</sub> > 0 such that for all *t* ≥ *t*<sub>*i*,*j*</sub>, *i*  $\int_0^{t} \int_0^{t}$  *j*. To establish the other implication, we will show that  $C^{\mu}(\infty) = \mathscr{A}^{\mu}$ , where  $\mathscr{A}^{\mu}$  is defined as follows:

$$
i \stackrel{\mathscr{A}^{\mu}}{\sim} j
$$
 if and only if  $\hat{W}^{\lambda}(J_i^{\mu}) = \hat{W}^{\lambda}(J_j^{\mu})$ .

Clearly all blocks of  $C^{\mu}(\infty)$  are sub-blocks of  $\mathscr{A}^{\mu}$ . It is thus sufficient to show that the blocks of  $\mathscr{A}^{\mu}$  have lengths of the same law as those of  $C^{\mu}(\infty)$ . Recall  $\kappa_{\lambda}$ . By [14, Lemma 5.8], blocks sizes of  $\mathscr{A}^{\mu}$  are i.i.d. with generating function given by,

$$
\mathbb{E}[z^{\#\mathscr{A}_1^{\mu}}] = 1 - \frac{\kappa_{\lambda}(\mu(1-z))}{\kappa_{\lambda}(\mu)}, \text{ for all } z \in (0, 1).
$$

Simple calculations from the expression of  $\kappa_{\lambda}$ , see (3.14), entail that  $\mathbb{E}[z^{\# \mathscr{A}_{1}}] = 1$  $e^{-\Psi^{\tau}(0+)\int_{\mu(1-z)}^{\mu}\frac{du}{\Psi(u)}}$ . Applying [14, Proposition 5.18], we see that  $\mathbb{E}[z^{\#\mathscr{A}_1^{\mu}}] = \mathbb{E}[z^{\#C_1^{\mu}(\infty)}]$  for any  $z \in (0, 1)$ .

It remains to deduce that  $\mathscr A$  is indeed the ancestral partition of the whole population. If  $x \stackrel{\mathscr{A}}{\sim} y$  and  $x \neq y$ , then there exists  $u > 0$ , such that  $x, y \in (W^{\lambda}(u-), W^{\lambda}(u))$ . By density of the set M, one can find for some large enough  $\mu$ , atoms  $J_x^{\mu}$ ,  $J_y^{\mu} \in \mathcal{M}$  such that

$$
W^{\lambda}(u-) < J_x^{\mu} < x \text{ and } W^{\lambda}(u) > J_y^{\mu} > y.
$$

By construction, the individuals  $J_x^{\mu}$  and  $J_y^{\mu}$  belong to the same interval of  $\mathscr{A}$ . We have seen above that it entails that almost surely for large enough  $t \geq 0$ ,  $\hat{X}_t(J_x^\mu) = \hat{X}_t(J_y^\mu)$ . Hence all individuals in the interval  $(J_x^{\mu}, J_y^{\mu})$  have a common ancestor, including *x* and *y*. This achieves the proof.  $\square$ 

**Proof of Proposition 3.8.** Recall the statement of Proposition 3.8, where  $\mathscr S$  denotes the support of the random measure  $d\hat{W}^{\lambda}$ . By Lemma 4.12,  $\hat{W}^{\lambda} = c_{\lambda,\lambda'}\hat{W}^{\lambda'}$  almost surely, this entails that  $\mathscr S$  does not depend on the parameter  $\lambda$ . Moreover, we see by Lemma 4.14 that  $\mathscr S$ is the range of a subordinator  $W^{\lambda}$  with Laplace exponent  $\kappa_{\lambda}$ . Having the Laplace exponent of  $W<sup>\lambda</sup>$  at hand, one can directly apply known results on the geometry of the range of a subordinator. By [3, Corollary 5.3], for all  $x > 0$ , almost surely dim<sub>*H*</sub>( $\mathcal{S} \cap [0, x]$ ) = ind( $\kappa_{\lambda}$ ) with  $\underline{\text{ind}}(\kappa_{\lambda}) := \liminf_{q \to \infty} \frac{\log \kappa_{\lambda}(q)}{\log q}$ . Recall  $\kappa_{\lambda}$ , one gets

$$
\underline{\text{ind}}(\kappa_{\lambda}) = \liminf_{q \to \infty} \frac{\Psi'(0+)}{\log q} \int_{\lambda}^{q} \frac{\mathrm{d}u}{\Psi(u)}.
$$

If  $\Psi'(\infty) := \lim_{u \to \infty} \frac{\Psi(u)}{u} < \infty$ , then we see that  $\int_{\lambda}^{q} \frac{du}{\Psi(u)} \sim_{\infty}$  $\frac{1}{\Psi'(\infty)} \log q$  and the result is established. If now  $\Psi'(\infty) = \infty$ , then for any large *D* > 0, for large *q*,

$$
\int_{\lambda}^{q} \frac{\mathrm{d} u}{\Psi(u)} \leq C_{\lambda} + \frac{\log q}{D},
$$

for some constant  $C_\lambda$ . We see therefore that  $\underline{\text{ind}}(\kappa_\lambda) \leq \frac{\Psi'(0+)}{D}$ . Since *D* is arbitrarily large, one can conclude that  $\dim_H(\mathscr{S} \cap [0, x]) = \frac{\Psi'(0+)}{\Psi'(\infty)} = 0$  a.s.  $\square$ 

We now establish Proposition 3.11 and look for the density of  $\hat{W}^{\lambda}(x)$  for fixed *x*.

**Proof of Proposition 3.11.** Let  $\lambda > 0$  and  $\theta > 0$ . Recall  $\kappa_{\lambda}$  and denote by  $e_{\theta}$ ,  $e_{\theta}$ two independent exponential random variables with parameter  $\theta$  and  $q$  respectively. By

Lemma 4.14, we have that  $\mathbb{P}(W^{\lambda}(e_{\theta}) \ge e_q) = \mathbb{P}(e_{\theta} \ge \hat{W}^{\lambda}(e_q))$ , or equivalently

$$
\mathbb{E}[e^{-\theta \hat{W}^{\lambda}(e_q)}] = 1 - \mathbb{E}[e^{-qW^{\lambda}(e_\theta)}].
$$

We deduce that

$$
\int_0^\infty \mathbb{E}[e^{-\theta \hat{W}^{\lambda}(x)}]e^{-qx}dx = \frac{1}{q} \left(1 - \mathbb{E}[e^{-\kappa_{\lambda}(q)\mathbf{e}_{\theta}}]\right) = \frac{1}{q} \left(1 - \frac{\theta}{\kappa_{\lambda}(q) + \theta}\right)
$$

$$
= \frac{1}{\theta} \frac{\kappa_{\lambda}(q)}{q} \frac{\theta}{\kappa_{\lambda}(q) + \theta}
$$

$$
= \frac{1}{\theta} \frac{\kappa_{\lambda}(q)}{q} \mathbb{E}[e^{-\kappa_{\lambda}(q)\mathbf{e}_{\theta}}] = \frac{1}{\theta} \int_0^\infty e^{-qu} \bar{\nu}_{\lambda}(u)du
$$

$$
\times \int_0^\infty e^{-qz} \mathbb{P}(W^{\lambda}(\mathbf{e}_{\theta}) \in \mathbf{d}_{z}),
$$

where we have used that

$$
\mathbb{E}[e^{-qW^{\lambda}(e_{\theta})}]=\mathbb{E}[e^{-\kappa_{\lambda}(q)e_{\theta}}] \text{ and } \frac{\kappa_{\lambda}(q)}{q}=\int_0^{\infty}e^{-qu}\bar{\nu}_{\lambda}(u)\mathrm{d}u.
$$

By the change of variable  $x = u + z$ , we obtain

$$
\int_0^\infty \mathbb{E}[e^{-\theta \hat{W}^\lambda(x)}]e^{-qx}dx = \frac{1}{\theta}\int_0^\infty e^{-qx}\int_0^\infty \bar{\nu}_\lambda(x-z)\mathbb{P}(W^\lambda(\mathbb{e}_\lambda) \in dz),
$$

and deduce that

$$
\mathbb{E}[e^{-\theta \hat{W}^{\lambda}(x)}] = \frac{1}{\theta} \int_0^x \bar{v}_{\lambda}(x-z) \mathbb{P}(W^{\lambda}(\mathbb{e}_{\theta}) \in dz)
$$
  
= 
$$
\int_0^{\infty} e^{-\theta u} \int_0^x \bar{v}_{\lambda}(x-z) \mathbb{P}(W^{\lambda}(u) \in dz) du.
$$

Thus, the density of  $\hat{W}^{\lambda}(x)$  is  $g^{\lambda}(u) := \int_0^x \bar{\nu}_{\lambda}(x-z) \mathbb{P}(W^{\lambda}(u) \in dz)$ . In the case  $\int_0^{\infty} \frac{du}{\Psi(u)} < \infty$ ,  $(W^{\infty}(u), u \ge 0)$  is a compound Poisson process with intensity  $v_{\infty}$ , and since  $\bar{v}_{\infty}(0) = 1$ , the formula (3.17) can be plainly checked.  $\square$ 

#### Declaration of competing interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

#### Acknowledgement

C.F's research is partially supported by LABEX MME-DII, Germany (ANR11-LBX-0023- 01).

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