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Near-Optimal Collaborative Learning in Bandits

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Abstract

This paper introduces a general multi-agent bandit model in which each agent is facing a finite set of arms and may communicate with other agents through a central controller in order to identify –in pure exploration– or play –in regret minimization– its optimal arm. The twist is that the optimal arm for each agent is the arm with largest expected *mixed* reward, where the mixed reward of an arm is a weighted sum of the rewards of this arm for *all agents*. This makes communication between agents often necessary. This general setting allows to recover and extend several recent models for collaborative bandit learning, including the recently proposed federated learning with personalization [30]. In this paper, we provide new lower bounds on the sample complexity of pure exploration and on the regret. We then propose a near-optimal algorithm for pure exploration. This algorithm is based on phased elimination with two novel ingredients: a data-dependent sampling scheme within each phase, aimed at matching a relaxation of the lower bound.

1 Introduction

Collaborative learning is a general machine learning paradigm in which a group of agents collectively train a learning algorithm. Some recent works have investigated how agents can efficiently perform sequential decision making in a collaborative context [36]. In particular, Shi et al. propose an interesting setting to tackle collaborative bandit learning when some level of personalization is required [30]. Personalization leads to the twist that each agent should play the best arm in a *mixed model* which is obtained as a combination of her *local model* with the local model of other agents. In this work, we introduce a more general model retaining this idea, that we call the weighted collaborative bandit model.

In this model, there are M agents and a finite number of K arms. When agent m samples arm k at time t , she gets to observe a *local reward* $X_{k,m}(t)$, which is drawn from a 1-sub-Gaussian¹ distribution of mean $\mu_{k,m}$, independently from past observations and from other agents' observations. However, this agent does not necessarily seek to maximize her local reward, but rather some notion of *mixed reward*, related to the utility of that arm for other agents. More specifically, we assume that agents share a *known* weight matrix $W := (w_{n,m})_{n,m} \in [0, 1]^{M \times M}$, such that $\sum_{n \in [M]} w_{n,m} = 1$

¹A random variable X is said to be σ^2 -sub-Gaussian if, for any $\lambda \in \mathbb{R}$, $\ln(\mathbb{E}[e^{\lambda(X - \mathbb{E}[X])}]) \leq \sigma^2 \lambda^2 / 2$.

for all $m \in [M]$, where $[n] := \{1, 2, \dots, n\}$. The mixed reward at time t for agent m and arm k is defined as a weighted average of the local rewards across all agents $X'_{k,m}(t) := \sum_{n=1}^M w_{n,m} X_{k,n}(t)$, and its expectation, called the expected mixed reward, is

$$\mu'_{k,m} := \sum_{n=1}^M w_{n,m} \mu_{k,n}.$$

We denote by $k_m^* := \arg \max_{k \in [K]} \mu'_{k,m}$ the arm with largest expected mixed reward for agent m , assumed unique. Besides the degenerated case in which $w_{n,m} = \mathbb{1}(n = m)$, in which each agent is solving their own bandit problem in isolation, the agents need to *communicate*; *i.e.*, to share information about their local rewards to other agents for everyone to be able to estimate their expected mixed rewards. A strategy for an agent is defined as follows: at each time t , each agent m samples an arm $\pi_m(t)$, based on the available information, and then observes a noisy local reward from this arm. She has also the option to communicate information (*e.g.*, empirical means of past local observations) to a central controller (or server), which will broadcast this information to all other agents. Just like in any multi-armed bandit model, several objectives may be considered, either related to maximizing (mixed) rewards –or, equivalently, minimizing regret– or identifying the best arms, while maintaining a reasonably small communication cost.

The weighted collaborative bandit model encompasses different frameworks previously studied in the literature. Notably, the paper [30] studies a special case in which, given a level of personalization $\alpha \in [0, 1]$, the mixed reward is an interpolation between $\mu_{k,m}$ and the average of local rewards $\frac{1}{M} \sum_{n=1}^M \mu_{k,n}$ which amounts to choosing $w_{m,m} = \alpha + \frac{1-\alpha}{M}$ and $w_{n,m} = \frac{1-\alpha}{M}$ for $n \neq m$. The authors consider the objective of minimizing the regret while minimizing for the number of communication rounds, under the name federated multi-armed bandit with personalization². In this paper, we mainly focus on the counterpart pure exploration problem in which agents should collaboratively identify their *own* optimal arm in terms of expected mixed reward, with high confidence, and using as few exploration rounds as possible. This extends the well-studied fixed-confidence best arm identification problem [12, 14, 20] to the weighted collaborative bandit setting.

Another related setting is collaborative pure exploration [18, 32, 6, 35], which considers M agents solving the *same* best arm identification problem. Most of these papers propose algorithms to solve this problem, while [32] and [21] also prove lower bounds on sample complexity for the fixed-confidence and the fixed-budget best arm identification. Unlike our framework, [6] consider asynchronous agents, which can only sample at some times, whereas [35, 36] consider agents which can only communicate to some of the other agents. The goal of collaborative pure exploration is to reduce sample complexity at the cost of some communication rounds. Our model recovers the synchronous setting when considering $\mu_{k,m} = \mu_{k,m'}$ for all arm k and agents m, m' and $W = Id_M$.

Besides collaborative learning, the work of [34] considers a similar weighted model but in a different, contextual bandit setting in which a central controller chooses in each round an arm for the unique agent (corresponding to a sub-population) that arrives, and aim at minimizing regret. Finally, the paper [29] considers a pure exploration task in which the value of an arm is the weighted average of its utility for M distinct populations (agents). In this setting $w_{n,m} = \alpha_n$, so that all agents have a common best arm, but the proposed algorithms do not aim at a low communication cost.

Collaborative learning in our general weighted model is also interesting beyond these examples. In particular, the work of [30] mentioned possible applications to recommendation systems, for which one may want to go beyond uniformly personalized learning. The “personalization” part means that we favour the local rewards of agent m over other agents’ observations in the identification of optimal arm k_m^* . But the introduction of a general weight matrix W in our framework allows *any* agent m to consider *any* linear combination of the other agents’ observations, and then, different degrees of personalization across agents.

Such a setting could also be appropriate for adaptive clinical trials on K therapies, run by M teams who have access to different sub-populations of patients. In this context, each sub-population is typically aiming at finding their (local) best treatment. However, solving their best arm identification in isolation may have a large sample complexity. If one is willing to assume that we have a

²Their setting, as well as ours, is neglecting some challenges typically addressed in (centralized) federated learning, such as privacy issues or dealing with communication interruption, this is why we prefer naming our framework “collaborative learning”.

weight matrix W for which the best mixed arm of each agent coincides with its local best arm, i.e. $k_m^* = \arg \max_{k \in [K]} \mu_{k,m}$ for any agent m , then solving best arm identification in the weighted collaborative bandit could have a much smaller sample complexity due to the sharing of information, while allowing the different clinical centers to communicate only once in a while. A first possibility to build such a weight matrix is to rely on a clustering of the different sub-populations so that the $\mu_{k,m}$ is supposed to be close to $\mu_{k,m'}$ (but not necessarily equal) for all k and all agents m, m' in the same cluster. Denoting by C_m the cluster to which agent m belongs, by setting $w_{m,n} = \mathbb{1}(C_m = C_n)/|C_m|$ we would have the mixed mean of each agent be very close to their local means. Another possibility is to rely on a similarity function S between sub-populations and define for all n, m $w_{m,n}$ to be proportional to $S(m, n)$. This similarity function could be obtained prior to the learning phase by computing the similarities between subpopulation biomarkers, for instance.

Related work Collaborative bandit learning has recently sparked wide interest in the multi-armed bandit literature. While some works do not deal with personalization [11, 31, 26], others have for instance studied the integration of arm features in a modified setting with personalization [19]. An interesting kernelized collaborative pure exploration problem was recently studied by [10]. In their model, both agents and arms are described by feature vectors, and there is a known kernel encoding the similarity between the mean reward of each (agent,arm) pair. This independent work follows a similar approach as ours and also propose a near-optimal phased elimination algorithm inspired by a lower bound, but the models and related lower bounds are significantly different.

In bandits working in collaboration, the need for a small communication cost indeed makes algorithms based on *phased eliminations* appealing. In such algorithms, agent(s) maintain a set of a *active* arms that are candidate for being optimal, and potentially eliminate arms from this set at the end of each sampling phase. Adaptivity to the observed rewards (and, in our case, communication) is only needed between sampling phases, which are typically long. This type of structure has been used in various bandit settings, both for regret minimization or pure exploration objectives [2, 18, 5, 13, 30, 3]. In some of these algorithms, including the one in [30] which motivates this paper, the number of samples gathered from an arm which is active in some phase r is fixed in advance. We believe that going beyond such a deterministic sampling scheme is crucial to achieve optimal performance with phased algorithms. In order to achieve (near)-optimality, other phased algorithms rely on computing an oracle allocation from the optimization problem associated with a lower bound on the sample complexity [13, 10]. In these works, based on this allocation, a total number of samples to collect in the current phase is computed, which depends on the identity of the surviving arms. The distribution of samples across arms for the current round is proportional to the oracle allocation, and is obtained through a rounding procedure. Compared to these works, our algorithm will show three distinctive features: an allocation inspired by a *relaxation* of the lower bound, which does not only depend on the identity of surviving arms, and an alternative to the rounding procedure.

Contributions The authors of [30] exhibit an algorithm using phased elimination for federated bandit learning with personalization, and prove a logarithmic regret bound, whose dependency in the parameters of the problem is conjectured to be sub-optimal. They also propose a heuristic improvement, based on a more adaptive exploration within each phase. In this work, we take a step further in identifying the problem-dependent complexity of bandit learning in our novel weighted collaborative bandit model –which includes the setting of [30] as a special case– both from the pure exploration and the regret perspective. We propose new information theoretic lower bounds, and a recipe to design algorithms (nearly) matching those for pure exploration. Our main algorithmic contribution is for weighted collaborative best arm identification. Our phased elimination-based algorithm achieves minimal exploration cost up to some logarithmic multiplicative factors, while using a constant amount of communication rounds. It relies on a novel data-dependent sampling scheme, which renders its analysis trickier. The structure of the algorithm can easily be extended to Top- N identification, that is, where each agent has to identify her own N best arms (instead of her best for $N = 1$) with respect to mixed rewards. We further compare our novel regret lower bound to what was conjectured in [30] for the particular case of federated learning with personalization.

Notation For both best arm identification and regret minimization, our complexity terms feature the (mixed) gaps of each agent m and arm k , defined by

$$\Delta'_{k,m} := \begin{cases} \mu'_{k_m^*,m} - \mu'_{k,m} & \text{if } k \neq k_m^* , \\ \min_{k \neq k_m^*} \Delta'_{k,m} & \text{otherwise .} \end{cases}$$

2 Collaborative Best-Arm Identification

The goal of collaborative best-arm identification is that each agent m identifies its optimal arm k_m^* by sampling the arms as little as possible and with few communication rounds. Formally, a collaborative Best Arm Identification (BAI) algorithm consists of a sampling rule π_m for each agent m , such that, at time t , either arm $\pi_m(t) \in [K]$ is sampled by m , or $\pi_m(t) = 0$; in that case, instead of picking an arm at time t , we allow the agent to remain idle, and not to select an arm.³ Similarly to the communication model studied in [32, 30], communication only happens at the end of local sampling rounds for all agents, when all agents are idle at the same time. Besides the sampling rule, the BAI algorithm uses a stopping rule τ which determines when exploration is over *for all agents*. The end of exploration is decided by the central server. Then, at time τ , each player outputs a guess for its optimal arm with respect to mixed rewards, denoted by \hat{k}_m .

Our goal is to construct a δ -correct strategy $\mathcal{A} = (\pi, \tau, \hat{k})$, which satisfies, for any model $\mu \in \mathbb{R}^{K \times M}$,

$$\mathbb{P}_\mu^{\mathcal{A}} (\forall m \in [M], \hat{k}_m = k_m^*) \geq 1 - \delta,$$

while achieving a small *exploration cost* (e.g. in high probability or in expectation)

$$\text{Exp}_\mu(\mathcal{A}) := \sum_{m=1}^M \sum_{k=1}^K N_{k,m}(\tau),$$

where $N_{k,m}(t) := \sum_{s=1}^t \mathbb{1}_{(\pi_m(s)=k)}$ is the number of selections of arm k by agent m up to time t , and a small *communication cost*, defined as

$$\text{Com}_\mu(\mathcal{A}) := \sum_{t=1}^{\tau} \mathbb{1}(\mathcal{I}_t),$$

where \mathcal{I}_t is the event that some information is shared between agents at round t .

In our setting, we do not put constraints on the type of information that is exchanged in each communication round—which can be interesting when we consider privacy issues [11, 36]—nor on the lengths of the messages. Each communication round has a unit cost. In a communication round, all agents send messages to the central server (e.g., estimates of their local means) and the server can send back arbitrary quantities or instructions (e.g., how many times each arm should be sampled in the next exploration phase, and when to communicate next).

Moreover, contrary to the works of [18, 32] on collaborative learning, we do not look at strategies explicitly minimizing for the number of communication rounds. Instead, our approach consists in proving a lower bound on the smallest possible exploration cost of a δ -correct algorithm which would communicate at every round; and then, finding an algorithm for which exploration cost matches this lower bound, while suffering a reasonable communication cost.

3 Lower Bound

We prove the following lower bound on the exploration cost of an algorithm in which all agents communicate to the central server their latest observation as soon as they received it. It holds for Gaussian rewards with variance $\sigma^2 = 1$, meaning that the reward from an arm k observed at time t by agent m will be $X_{k,m}(t) = \mu_k + \varepsilon_t$, where $\varepsilon_t \sim \mathcal{N}(0, 1)$. We further assume that the weight matrix W satisfies $w_{m,m} \neq 0$ for any agent $m \in [M]$.

Theorem 1. *Let μ be a fixed matrix of means in $\mathbb{R}^{K \times M}$. For any $\delta \in (0, 1/2]$, let \mathcal{A} be a δ -correct algorithm under which each agent communicates each reward to the central server after it is observed, and let us denote for any $k \in [K]$, $m \in [M]$, $\tau_{k,m} := \mathbb{E}_\mu^{\mathcal{A}} [N_{k,m}(\tau)]$, where τ is the stopping time. For any $m \in [M]$ and $k \neq k_m^*$, it holds that*

$$\sum_n w_{n,m}^2 \left(\frac{1}{\tau_{k,n}} + \frac{1}{\tau_{k_m^*,n}} \right) \leq \frac{(\Delta'_{k,m})^2 / 2}{\log(1/(2.4\delta))},$$

and therefore $\mathbb{E}_\mu [\text{Exp}_\mu(\mathcal{A})] \geq T_W^*(\mu') \log\left(\frac{1}{2.4\delta}\right)$, where

³Note that, in a regret setting, an idle agent may exploit its empirical best arm.

$$T_W^*(\mu) := \min_{t \in (\mathbb{R}^+)^{K \times M}} \left\{ \sum_{(k,m) \in [K] \times [M]} t_{k,m} : \forall m, k \neq k_m^*, \sum_{n \in [M]} w_{n,m}^2 \left(\frac{1}{t_{k,n}} + \frac{1}{t_{k_m^*,n}} \right) \leq \frac{(\Delta'_{k,m})^2}{2} \right\}.$$

The proof, given in Appendix A.1, uses standard change-of-distribution arguments, together with classical results from constrained optimization. Note that, for $M = 1$, we recover the complexity of best arm identification in a Gaussian bandit model [15].

Computing the complexity term The optimization problem which defines $T_W^*(\mu)$ belongs to the family of disciplined convex optimization problems, and can be numerically solved using available solvers, such as CVXPY [1, 9]. We now illustrate, on a small example, the possible reduction in exploration cost that can be obtained by solving weighted collaborative best arm identification instead of M parallel best arm identification problems. We consider $K = M = 2$ and a similarity $S(1, 2) = 0.9$ between the two agents, which yields the following normalized weight matrix

$$W = \frac{1}{1.9} \begin{bmatrix} 1 & 0.9 \\ 0.9 & 1 \end{bmatrix}.$$

Considering the following matrix of expected rewards

$$\mu = \begin{bmatrix} 0.9 & 0.8 \\ 0.1 & 0.5 \end{bmatrix},$$

for which arm 1 is the local best arm for both agents and is also their best mixed arm, we obtain $T_W^*(\mu) \approx 28$. However, if each agent solves its own best arm identification problem in isolation (which amounts to using $W = \text{Id}_2$), the resulting exploration cost scales with $T_{\text{Id}_2}^*(\mu) \approx 101 > 3T_W^*(\mu)$.

From lower bounds to algorithms In single-agent pure exploration tasks, lower bounds are usually guidelines to design optimal algorithms, as they allow to recover an oracle allocation (*i.e.*, the arg min for $t \in (\mathbb{R}^+)^{K, M}$ in the definition of $T_W^*(\mu)$) which algorithms can try to achieve by using some tracking [15, 10, 29]. Yet, these approaches may be computationally expensive, as they solve the optimization problem featured in the lower bound in every round.

In the next section, we will propose an alternative approach for our collaborative setting, which exploits the knowledge of the lower bound within a phased elimination algorithm. This is crucial to maintain a small communication cost and also permit to reduce the computational complexity compared to a pure tracking approach. Our algorithm will rely on a *relaxed* complexity term $\tilde{T}_W^*(\mu)$, which is within constant factors of $T_W^*(\mu)$, as proved in Appendix C.1.

Lemma 1. *Introducing the quantity*

$$\tilde{T}_W^*(\mu) := \min_{t \in (\mathbb{R}^+)^{K \times M}} \left\{ \sum_{(k,m) \in [K] \times [M]} t_{k,m} : \forall m, \forall k \in [K], \sum_{n \in [M]} \frac{w_{n,m}^2}{t_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2} \right\},$$

it holds that $\tilde{T}_W^*(\mu) \leq T_W^*(\mu) \leq 2\tilde{T}_W^*(\mu)$.

Compared to $T_W^*(\mu)$, a nice feature of $\tilde{T}_W^*(\mu)$ is that its constraint set does not depend on the knowledge of $(k_m^*)_{m \in [M]}$, which will allow us to design algorithms that do not suffer too much from bad empirical guesses for k_m^* in early phases. We further remark that the computation of $\tilde{T}_W^*(\mu)$ and that of its associated oracle allocation (see Definition 1) are slightly easier than for $T_W^*(\mu)$. Indeed, the optimization problem which defines $\tilde{T}_W^*(\mu)$ can be decoupled across arms. Computing for every arm $k \in [K]$ the vector

$$\tilde{\tau}^k = \underset{\tau^k \in (\mathbb{R}^+)^M}{\text{argmin}} \sum_{(k,m) \in [K] \times [M]} \tau_m^k \text{ s.t. } \forall m, \sum_{n \in [M]} \frac{w_{n,m}^2}{\tau_n^k} \leq \frac{(\Delta'_{k,m})^2}{2},$$

we obtain the argmin in $\tilde{T}_W^*(\mu)$ by setting $(t_{k,m})_{k,m} = (\tilde{\tau}_m^k)_{k,m}$. The computation of $\tilde{T}_W^*(\mu)$ can therefore be done by solving K disciplined optimization problems (*e.g.*, with CVXPY [1, 9]) involving M variables, instead of one optimization problem with $K \times M$ variables.

Definition 1. For any $\Delta \in (\mathbb{R}^+)^{K \times M}$, the oracle $\tilde{\mathcal{P}}^*(\Delta)$ is

$$\arg \min_{(\tau_{k,m})_{k,m} \in (\mathbb{R}^+)^{K \times M}} \sum_{k,m} \tau_{k,m} \text{ s.t. } \forall m \in [M], \forall k \in [K], \sum_{n \in [M]} \frac{w_{n,m}^2}{\tau_{k,n}} \leq \frac{(\Delta_{k,m})^2}{2}.$$

With this notation, observe that $\tilde{T}^*(\mu') = \sum_{k,m} \tau_{k,m}$, where $(\tau_{k,m})_{k,m} \in \tilde{\mathcal{P}}^*(\Delta')$. The following lemma will be useful to compare values from different oracle problems. The full lemma along with its proof is available in Appendix C.1.

Lemma 2. Consider $\Delta, \Delta' \in (\mathbb{R}^+)^{K \times M}$, such that $\tau \in \tilde{\mathcal{P}}^*(\Delta)$ and $\tau' \in \tilde{\mathcal{P}}^*(\Delta')$. Moreover, assume that there is a positive constant β such that: $\forall k, m, \Delta'_{k,m} \leq \beta \Delta_{k,m}$. Then

$$\frac{1}{\beta^2} \sum_{k,m} \tau_{k,m} \leq \sum_{k,m} \tau'_{k,m}.$$

4 A Near-Optimal Algorithm For Best Arm Identification

We now introduce an algorithm for collaborative best-arm identification, called W-CPE-BAI for Weighted Collaborative Phased Elimination, stated as Algorithm 1. To present an analysis of this algorithm, we assume that, for any $k, m \in [K] \times [M]$, $\mu_{k,m} \in [0, 1]$, and that local rewards $(X_{k,m}(t))_{k,m,t}$ are 1-sub-Gaussian.

W-CPE-BAI proceeds in phases, indexed by r . In phase r , we let $B_m(r)$ be the set of active arms for agent m , and $B(r) = \cup_m B_m(r)$ be the set of arms that are active for at least one agent. The algorithm maintains proxies for the gaps $(\tilde{\Delta}_{k,m}(r))_{k \in [K], m \in [M]}$ that are halved at the end of each phase for arms that remain active. At the beginning of each round, the oracle allocation $t(r)$, with respect to the proxy gaps, is computed, as well as the number of new samples $d_{k,m}(r)$ that player m should get from arm k in phase r . $d_{k,m}(r)$ is defined such that the total number of selections of arm k by agent m becomes close to (a quantity slightly larger than) $t_{k,m}(r) \log(1/\delta)$. We observe that any arm $k \notin B(r)$ will not get any new samples in phase r , as the proxy gaps $(\tilde{\Delta}_{k,n}(r))_n$ are identical to those in the previous phase; therefore $t_{k,n}(r) = t_{k,n}(r-1)$. In contrast to prior works, where the allocation in each round only depends on the identity of the surviving arms and the round index [13, 10], in W-CPE-BAI it also depends on when the arms have been eliminated (which condition the value of their frozen proxy gaps).

After each agent m samples arm k $d_{k,m}(r)$ times, they send their local means $\hat{\mu}_{k,m}(r)$ to the central server, which computes the mixed mean estimates $\hat{\mu}'_{k,n}(r) := \sum_{m=1}^M w_{n,m} \hat{\mu}_{k,m}(r)$. The active sets $(B_m(r))_{m \in [M]}$ of all agents are then updated by removing arms whose mixed means are too small. As in several prior works [22, 30], we rely on confidence intervals to perform these eliminations. However, constructing confidence intervals on the mixed means—which are linear combinations of the local means—under our adaptive sampling rule is more challenging than when the number of samples from an active arm in phase r is fixed in advance—which is the case for instance in the algorithm in [30]. The width of our confidence intervals scales with the following quantity

Definition 2. For any k, m , and round $r \geq 0$, we define

$$\Omega_{k,m}(r) := \sqrt{\beta_\delta(n_{k,\cdot}(r)) \sum_{n=1}^M \frac{w_{n,m}^2}{n_{k,n}(r)}},$$

where $n_{k,m}(r)$ is the number of times arm k was selected by agent m by the end of phase r (included), and $N \mapsto \beta^\delta(N)$ is a threshold function defined for any $N \in (\mathbb{R}^+)^M$.

Leveraging some recent time-uniform concentration inequalities [23], we exhibit below a choice of threshold that yields valid confidence intervals on the mixed means (by "projecting" confidence intervals that can be obtained on local means, see Proposition 24 in [23]). The fact that the confidence interval depends on the random number of past draws (and not just the index of the round) leads to some non-trivial complication in the analysis, with the introduction of quantities $(d_{k,m})_{k,m}$. The proof is given in Appendix C.2, where we also provide an explicit expression of the function g_M .

Algorithm 1 Weighted Collaborative Phased Elimination for Best Arm Identification (W-CPE-BAI)

Input: $\delta \in (0, 1)$, M agents, K arms, weights matrix W

Initialize $r \leftarrow 0$, $\forall k, m, \tilde{\Delta}_{k,m}(0) \leftarrow 1, n_{k,m}(0) \leftarrow 1, \forall m, B_m(0) \leftarrow [K]$

Draw each arm k by each agent m once

repeat

 # Central server

$B(r) \leftarrow \cup_{m \in [M]} B_m(r)$

 Compute $t(r) \leftarrow \tilde{\mathcal{P}}^* \left(\left(\sqrt{2} \tilde{\Delta}_{k,m}(r) \right)_{k,m} \right)$

 For any $k \in [K]$, compute

$$(d_{k,m}(r))_{m \in [M]} \leftarrow \arg \min_{d \in \mathbb{N}^M} \sum_m d_m \text{ s.t. } \forall m \in [M], \frac{n_{k,m}(r-1) + d_m}{\beta_\delta(n_{k,\cdot}(r-1) + d)} \geq t_{k,m}(r)$$

 Send to each agent m $(d_{k,m}(r))_{k,m}$ and $d_{\max} := \max_{n \in [M]} \sum_{k \in [K]} d_{k,n}(r)$

 # Agent m

 Sample arm $k \in B(r)$ $d_{k,m}(r)$ times, so that $n_{k,m}(r) = n_{k,m}(r-1) + d_{k,m}(r)$

 Remain idle for $d_{\max} - \sum_{k \in [K]} d_{k,m}(r)$ rounds

 Send to the server empirical mean $\hat{\mu}_{k,m}(r) := \sum_{s \leq n_{k,m}(r)} X_{k,m}(s) / n_{k,m}(r)$ for any $k \in [K]$

 # Central server

 Compute the empirical mixed means $(\hat{\mu}'_{k,m}(r))_{k,m}$ based on $(\hat{\mu}_{k,m}(r))_{k,m}$ and W

 // Update set of candidate best arms for each user

for $m = 1$ **to** M **do**

$$B_m(r+1) \leftarrow \left\{ k \in B_m(r) \mid \hat{\mu}'_{k,m}(r) + \Omega_{k,m}(r) \geq \max_{j \in B_m(r)} \left(\hat{\mu}'_{j,m}(r) - \Omega_{j,m}(r) \right) \right\}$$

end for

 // Update the gap estimates

 For all $k, m, \tilde{\Delta}_{k,m}(r+1) \leftarrow \tilde{\Delta}_{k,m}(r) \times (1/2)^{\mathbb{1}(k \in B_m(r+1) \wedge |B_m(r+1)| > 1)}$

$r \leftarrow r + 1$

until $\forall m \in [M], |B_m(r)| \leq 1$

Output: $\{k \in B_m(r) : m \in [M]\}$

Lemma 3. *Let us define*

$$\beta_\delta(N) := 2 \left(g_M \left(\frac{\delta}{KM} \right) + 2 \sum_{m=1}^M \ln(4 + \ln(N_m)) \right),$$

for any $N \in (\mathbb{N}^*)^M$, where g_M is some non-explicit function, defined in [23], that satisfies $g_M(\delta) \simeq \log(\frac{1}{\delta}) + M \log \log(\frac{1}{\delta})$. Then the good event

$$\mathcal{E} := \left\{ \forall r \in \mathbb{N}, \forall m, \forall k, \left| \hat{\mu}'_{k,m}(r) - \mu'_{k,m} \right| \leq \Omega_{k,m}(r) \right\}$$

holds with probability larger than $1 - \delta$.

From this lemma, it easily follows that W-CPE-BAI is δ -correct for the above choice of threshold function, as, for any agent m , no good arm k_m^* can ever be eliminated from $B_m(r)$ at round r . The fact that the sample complexity of W-CPE-BAI scales with $\tilde{T}^*(\mu')$ on the good event \mathcal{E} comes from the interplay between the expression of $\Omega_{k,m}(r)$ (which, up to the threshold function, is exactly one of the constraints featured in the lower bound) together with the definition of the allocation $t(r)$, which leads to the following crucial result in our analysis

Lemma 4. *On \mathcal{E} , $\forall k, m, r \geq 0, \Omega_{k,m}(r) \leq \tilde{\Delta}_{k,m}(r)$.*

Proof. For any round r and arm k , by Algorithm 1 and the definition of oracle $t_{k,\cdot}(r)$, for any agent m ,

$$\Omega_{k,m}(r) = \sqrt{\sum_n w_{n,m}^2 \frac{\beta_\delta(n_{k,\cdot}(r-1) + d_{k,\cdot})}{n_{k,n}(r-1) + d_{k,n}(r)}} \leq \begin{cases} \sqrt{\sum_n \frac{w_{n,m}^2}{t_{k,n}(r)}} \leq \tilde{\Delta}_{k,m}(r) & \text{if } k \in B(r), \\ \sqrt{\sum_n \frac{w_{n,m}^2}{t_{k,n}(r'_k)}} \leq \tilde{\Delta}_{k,m}(r'_k) = \tilde{\Delta}_{k,m}(r) & \text{otherwise,} \end{cases}$$

where when $k \notin B(r)$, $r'_k := \sup\{r' \geq 0 : k \in B(r')\}$, and we use the fact that $d_{k,m}(r) = 0$ when $k \notin B(r)$. \square

We did not put much emphasis on the way communications are performed between the agents and the central server, as several choices are possible. The important part is that the server receives all values of the local means $(\hat{\mu}_{k,m}(r))_{k,m}$ at the end of round r . Our suggestion is that the central server maintains the sets $(B_m(r))_{m \in [M]}$, calls the oracle, and sends to all agents their values of $(d_{k,m}(r+1))_{k,m}$ at the end of each phase r . In any case, the number of communication rounds in our definition will be equal to the number of phases used by W-CPE-BAI. All in all, we prove

Theorem 2. *With probability $1 - \delta$, W-CPE-BAI outputs the optimal arm for each agent with an exploration cost at most*

$$32\tilde{T}_W^*(\mu) \log_2(8/\Delta'_{\min}) \log\left(\frac{1}{\delta}\right) + o_\delta\left(\log\left(\frac{1}{\delta}\right)\right),$$

and at most $\lceil \log_2(8/\Delta'_{\min}) \rceil$ communication rounds, where $\Delta'_{\min} := \min_{k \in [K], m \in [M]} \Delta'_{k,m}$.

Proof sketch The detailed proof is given in Appendix B, where we also provide an explicit upper bound on the exploration cost. We let R denote the (random) number of phases used by the algorithm before stopping. On the good event \mathcal{E} , we can prove that the algorithm never eliminates k_m^* therefore $R := \max_m \max_{k \neq k_m^*} R_{k,m}$ where $R_{k,m}$ is the last phase in which $k \in B_m(r)$. Using Lemma 4, we can easily establish that

$$R_{k,m} \leq r_{k,m} := \min\{r \geq 0 : 4 \times 2^{-r} < \Delta'_{k,m}\}$$

which satisfies $r_{k,m} \leq \log_2(8/\Delta'_{k,m})$. This yields $R \leq \log_2(8/\Delta'_{\min})$, and further permits to prove that the proxy gaps can be lower bounded by the true gaps:

$$\forall r \leq R, \forall k \in [K], \forall m \in [M], \tilde{\Delta}_{k,m}(r) \geq \frac{1}{8} \Delta'_{k,m}.$$

See Corollary 2 in Appendix B. Using the monotonicity properties of the oracle that are stated in Lemma 2, we can then establish that the allocation $t(r)$ computed from the proxy gaps in the algorithm satisfies

$$\forall r \leq R, \sum_{k \in [K], m \in [M]} t_{k,m}(r) \leq 32\tilde{T}_W^*(\mu). \quad (1)$$

To upper bound the exploration cost, the next step is to relate $n_{k,m}(R)$ to the oracle allocations. To do so, we observe that if $R'_{k,m}$ is the last round before R such that $d_{k,m}(r) \neq 0$ (i.e. the last round in which arm k is actually sampled by agent m ; then $n_{k,m}(R) = n_{k,m}(R'_{k,m})$), we have by definition of the $(d_{k,m}(r))_{k,m,r}$ that

$$n_{k,m}(R'_{k,m}) \leq t_{k,m}(R'_{k,m})\beta_\delta(n_{k,\cdot}(R'_{k,m})) + 1 \leq t_{k,m}(R'_{k,m})\beta_\delta(n_{k,\cdot}(R)) + 1.$$

See Lemma 12 in Appendix B. We can then upper bound $\tau := \sum_{k,m} n_{k,m}(R)$ as follows

$$\tau \leq \sum_{k,m} t_{k,m}(r'_{k,m})\beta_\delta(n_{k,\cdot}(R)) + KM \leq \sum_{k,m} \sum_{r \leq R} t_{k,m}(r)\beta^*(\tau) + KM,$$

$$\text{and } \tau \leq R \times 32\tilde{T}_W^*(\mu)\beta^*(\tau) + KM, \text{ where we use (1) and introduce} \quad (2)$$

$$\beta^*(\tau) := 2 \left(g_M \left(\frac{\delta}{KM} \right) + 2M \ln(4 + \ln(\tau)) \right).$$

The end of the proof consists in using the known upper bound on R , and finding an upper bound for the largest τ satisfying the inequality in (2). \square

Discussion Theorem 2 proves that W-CPE-BAI is matching the exploration lower bound of Theorem 1 in a regime where δ is small, up to multiplicative constants, including a logarithmic term in $1/\Delta'_{\min}$. It achieves this using only $\lceil \log_2(8/\Delta'_{\min}) \rceil$ communication rounds. We note that a similar extra multiplicative logarithmic factor is present in the analysis of near-optimal phased algorithms in other contexts [13, 10]. Such a quantity appears as an upper bound on the number of phases, and may be a price to pay for the phased structure.

On the communication cost We argue that the communication cost of W-CPE-BAI is actually of the same order of magnitude as that featured in some related work. In [30], which is the closest setting to our framework, the equivalent number of communication rounds p needed to solve the regret minimization problem is upper bounded by $\mathcal{O}(2 \log_2(8/(\sqrt{M}\Delta'_{\min})))$. In the setting of collaborative learning—where M agents face the same set of arm distributions and $W = Id$ —[18] in their Theorem 4.1 prove that an improvement of multiplicative factor $1/M$ on the exploration cost for a traditional best arm identification algorithm can be reached by using at most $\lceil \log_2(1/\Delta_{\min}) \rceil$ communication rounds, where Δ_{\min} is the gap between the best and second best arms.

Experimental validation We propose in Appendix E an empirical evaluation of W-CPE-BAI for the weight matrix $w_{m,n} = \alpha \mathbb{1}(n = m) + \frac{1-\alpha}{M}$ which corresponds to the setting studied by [30]. In this particular case, we propose as a baseline a counterpart of the regret algorithm of [30] which we call PF-UCB-BAI. Our experiments on a synthetic instance show that W-CPE-BAI and PF-UCB-BAI have similar performances in terms of exploration cost and that W-CPE-BAI becomes better when the level of personalization α is smaller than 0.5. Moreover, W-CPE-BAI uses less rounds of communication than PF-UCB-BAI for all values of α . Finally, the near-optimality of W-CPE-BAI is empirically observed when compared to an oracle algorithm which has access to the true gaps. We refer the reader to Appendix E for further details on the optimization libraries that were used.

Remark 1. *The analysis and the structure of Algorithm 1 have the potential to be extended to other pure exploration problems, with similar guarantees. In Appendix D, we illustrate this claim by tackling Top- N identification.*

5 Regret Lower Bound

In contrast to the BAI setting, [30] considered the objective of minimizing the regret

$$\mathcal{R}_\mu(T) := \mathbb{E} \left[\sum_{m=1}^M \sum_{t=1}^T (\mu'_{k_m^*, m} - \mu_{\pi_m(t), m}) \right].$$

They provided a conjecture on the lower bound on regret in personalized federated learning [see, 30, Conjecture 1]. As mentioned in introduction, their reward model is a special case of ours with weights $w_{m,m} = \alpha + \frac{1-\alpha}{M}$ and $w_{n,m} = \frac{1-\alpha}{M}$ for $n \neq m$. In this section, we prove a regret lower bound with general weights that proves in particular Conjecture 1 in [30].

We prove the following result, for an algorithm that selects in each round exactly one arm for each agent, and all agents communicate after each round. This lower bound holds for Gaussian arms with variance $\sigma^2 = 1$.

Theorem 3. *Any uniformly efficient algorithm⁴ in which all agents communicate after each round satisfies*

$$\liminf_{T \rightarrow \infty} \frac{\mathcal{R}(T)}{\log(T)} \geq C_W^*(\mu),$$

where

$$C_W^*(\mu) := \min_{c=(c_{k,m})_{k,m:k_m^* \neq k}} \left\{ \sum_{k=1}^K \sum_{m:k_m^* \neq k} c_{k,m} \Delta'_{k,m} : \forall k \in [K], \forall m \in [M], \sum_{n:k_n^* \neq k} \frac{w_{n,m}^2}{c_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2} \right\}.$$

We recall that for any agent m , $\Delta'_{k_m^*, m} := \min_{k' \neq k_m^*} \Delta'_{k', m}$.

⁴A uniformly efficient algorithm satisfies $\mathcal{R}_\mu(T) = o(T^\gamma)$ for any $\gamma \in (0, 1)$ and any possible instance μ .

Theorem 3 may be viewed as an extension of the lower bound of [25] to the collaborative setting. Its proof, given in Appendix A.2, uses classical ingredients for regret lower bounds in (single agent) structured bandit models [7, 17].

The generic approaches proposed in [7, 8] for optimal regret minimization in structured bandits could therefore be useful. However, to turn them into a reasonable algorithm for the collaborative setting, we would need to preserve the phased elimination structure. Analogous to the relaxation in the BAI setting, we can also define a relaxed complexity term $\tilde{C}_M(\mu)$, which does not require the knowledge of the best arms, and is within constant factors of $C(\mu)$ by the following Lemma 5 (which proof is given in Appendix C.3).

Lemma 5. *Introducing the quantity*

$$\tilde{C}_W^*(\mu) := \min_{c \in (\mathbb{R}^+)^{K \times M}} \left\{ \sum_{k=1}^K \sum_{m=1}^M c_{k,m} \Delta'_{k,m} \text{ s.t. } \forall k \in [K], \forall m \in [M], \sum_{n=1}^M \frac{w_{n,m}^2}{c_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2} \right\},$$

it holds that $C_W^*(\mu) \leq \tilde{C}_W^*(\mu) \leq 4C_W^*(\mu)$.

The constrained optimization problems governing the regret lower bound are subtly different from those in the BAI setting. In regret minimization, unknown gaps $(\Delta'_{k,m})_{k,m}$ appear both in the constraints and in the objective. In contrast, in BAI, $(\Delta'_{k,m})_{k,m}$ only appear in the constraints. Due to this difference, designing a similar algorithm for regret minimization leads to an extra multiplicative $\mathcal{O}(1/\Delta'_{\min})$ factor in the upper bound. A very similar issue exists in a different structured bandit problem, bandits on graphs with side observations, where arms are connected through a graph with unweighted edges, and the agent receives the reward associated with the selected arm *and its neighbors* at a given round. In [4, Problem P1], authors describe a constrained linear optimization problem which governs the regret lower bound, and face the same issue of scaling in $\mathcal{O}(1/\Delta'_{\min})$. Dealing with this problem would be an interesting subsequent work.

6 Conclusion

This paper introduced a general framework for collaborative learning in multi-armed bandits. Our work presents two novel lower bounds: one on the exploration cost for pure exploration, and another on cumulative regret. The latter permits to prove a prior conjecture on the topic. Moreover, we propose a phased elimination algorithm for fixed-confidence best arm identification. This algorithm tracks the optimal allocation from the pure exploration lower bound through a novel approach solving a relaxed optimization problem linked to the lower bound. The exploration cost of this algorithm is matching the lower bound up to logarithmic factors. This strategy can be extended to other pure exploration problems, such as Top- N identification, as shown in Appendix D.

As mentioned in introduction, our collaborative setting was motivated by the design of collaborative adaptive clinical trials for personalized drug recommendations, where several patient subpopulations (for instance, representing several subtypes of cancer) are considered and sequentially treated. However, in practice, especially when dealing with patient data, disclosing the mean response values to the central server should be handled with care to preserve the anonymity of the patients. A possible solution to overcome this problem would be to carefully combine our algorithm with a data privacy-preserving method, for instance by adding some noise to the data [11].

Our code and run traces are available in an open-source repository.⁵

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⁵<https://github.com/clreda/near-optimal-federated>

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Checklist

1. For all authors...
 - (a) Do the main claims made in the abstract and introduction accurately reflect the paper’s contributions and scope? [Yes] , a general framework which encompasses several prior works is described in Section 1; a lower bound on the sample complexity for a pure exploration problem is derived in Section 3, whereas a lower bound on regret on the regret minimization counterpart of the problem is shown in Section 5. An algorithm for best arm identification, using a novel approach to achieve near-optimality, is described and analyzed in Section 4.
 - (b) Did you describe the limitations of your work? [Yes] , we explained that the weight matrix had to be known in advance (Section 1), although there is an application case where it can be computed. Moreover, we explained in conclusion to our work the issue faced when applying the same technique to regret minimization (Section 6).
 - (c) Did you discuss any potential negative societal impacts of your work? [Yes] , we remarked in our discussion (Section 6) that our collaborative algorithm might not be adapted to private, sensitive data, but might still be interesting for dealing with collaborative adaptive clinical trials between trusting teams.
 - (d) Have you read the ethics review guidelines and ensured that your paper conforms to them? [Yes]
2. If you are including theoretical results...
 - (a) Did you state the full set of assumptions of all theoretical results? [Yes] , the assumptions are mentioned both in the statement of the results (Section 3, Section 4, Section 5) and the proofs (Appendix).
 - (b) Did you include complete proofs of all theoretical results? [Yes] , proofs in full are included in Appendix.
3. If you ran experiments...
 - (a) Did you include the code, data, and instructions needed to reproduce the main experimental results (either in the supplemental material or as a URL)? [Yes] , the code and run traces are available in a open-source repository. ⁶
 - (b) Did you specify all the training details (e.g., data splits, hyperparameters, how they were chosen)? [Yes] , this information is reported in Appendix E.
 - (c) Did you report error bars (e.g., with respect to the random seed after running experiments multiple times)? [Yes] , tables in Appendix E reports the average values \pm their standard deviation.
 - (d) Did you include the total amount of compute and the type of resources used (e.g., type of GPUs, internal cluster, or cloud provider)? [Yes] , this information is provided in Appendix E (paragraph “Numerical considerations”).
4. If you are using existing assets (e.g., code, data, models) or curating/releasing new assets...
 - (a) If your work uses existing assets, did you cite the creators? [N/A]
 - (b) Did you mention the license of the assets? [N/A]
 - (c) Did you include any new assets either in the supplemental material or as a URL? [N/A]
 - (d) Did you discuss whether and how consent was obtained from people whose data you’re using/curating? [N/A]
 - (e) Did you discuss whether the data you are using/curating contains personally identifiable information or offensive content? [N/A]
5. If you used crowdsourcing or conducted research with human subjects...
 - (a) Did you include the full text of instructions given to participants and screenshots, if applicable? [N/A]
 - (b) Did you describe any potential participant risks, with links to Institutional Review Board (IRB) approvals, if applicable? [N/A]
 - (c) Did you include the estimated hourly wage paid to participants and the total amount spent on participant compensation? [N/A]

⁶<https://github.com/clreda/near-optimal-federated>

A Proof of the Lower Bounds

A.1 Lower Bound on the Exploration Cost: Proof of Theorem 1

Let us denote, for any model $\mu \in \mathbb{R}^{K \times M}$ and agent m , $k_m^* := \arg \max_{k \in [K]} \mu'_{k,m}$ (which is assumed unique). Define the set of alternative models in $\mathbb{R}^{K \times M}$ with respect to μ :

$$\text{Alt}(\mu) := \left\{ \lambda : \exists m, \exists k \neq k_m^* : \lambda'_{k,m} > \lambda'_{k_m^*,m} \right\},$$

where $\lambda'_{k,m} := \sum_{n \in [M]} w_{n,m} \lambda_{k,n}$ for any arm k and agent m . Assume that stopping time τ is almost surely finite under μ for algorithm \mathcal{A} . Let event $\mathcal{E}_\mu := \{ \exists m : \hat{k}_m \neq k_m^* \}$. As algorithm \mathcal{A} is δ -correct, it holds that $\mathbb{P}_\mu(\mathcal{E}_\mu) \leq \delta$ and $\mathbb{P}_\lambda(\mathcal{E}_\mu) \geq 1 - \delta$ for all $\lambda \in \text{Alt}(\mu)$. As this event belongs to the filtration generated by all past observations up to the final stopping time τ , using Theorem 1 from [15] and $\delta \leq 1/2$, it holds that

$$\frac{1}{2} \sum_{k,m} \tau_{k,m} (\mu_{k,m} - \lambda_{k,m})^2 \geq \log \left(\frac{1}{2.4\delta} \right). \quad (3)$$

We first prove that, for any k, m , $\tau_{k,m} > 0$. Indeed, if $w_{m,m} \neq 0$, it is possible to pick λ that only differs from μ by the entry $\lambda_{k,m}$, in such a way that arm k becomes optimal (or sub-optimal) for user m . From Equation (3), we get $\frac{1}{2} \tau_{k,m} (\mu_{k,m} - \lambda_{k,m})^2 > 0$ and the conclusion follows.

We now fix agent m and $k \neq k_m^*$, and try to find a more informative alternative model λ . We look for it in a family of alternative models under which only arms k and k_m^* are modified, for all agents, in order to make arm k optimal for agent m . Given two nonnegative sequences $(\delta_n)_{n \in [M]}$ and $(\delta'_n)_{n \in [M]}$, we define $\lambda = (\lambda_{k',n})_{k'}$ such that

$$\begin{cases} \lambda_{k',n} &= \mu_{k',n} \text{ if } k' \notin \{k, k_m^*\}, \\ \lambda_{k,n} &= \mu_{k,n} + \delta_n, \\ \lambda_{k_m^*,n} &= \mu_{k_m^*,n} - \delta'_n, \end{cases}$$

that satisfies

$$\sum_{n \in [M]} w_{n,m} (\delta_n + \delta'_n) \geq \Delta'_{k,m}. \quad (4)$$

Now arm k is optimal for agent m . From Equation (3),

$$\sum_n \tau_{k,n} \frac{\delta_n^2}{2} + \sum_n \tau_{k_m^*,n} \frac{(\delta'_n)^2}{2} \geq \log \left(\frac{1}{2.4\delta} \right).$$

Hence, it holds that

$$\inf_{\delta, \delta': (4) \text{ holds}} \left[\sum_n \tau_{k,n} \frac{\delta_n^2}{2} + \sum_n \tau_{k_m^*,n} \frac{(\delta'_n)^2}{2} \right].$$

The infimum can be computed in closed form using constraint optimization. Introducing a Lagrange multiplier λ , from the KKT conditions, we get that, for any agent n ,

$$\begin{aligned} \tau_{k,n} \delta_n - \lambda w_{n,m} &= 0, \\ \tau_{k_m^*,n} \delta'_n - \lambda w_{n,m} &= 0, \\ \lambda \left(\sum_{n'} w_{n',m} (\delta_{n'} + \delta'_{n'}) - \Delta'_{k,m} \right) &= 0. \end{aligned}$$

Using furthermore that $\tau_{k,n}$ and $\tau_{k_m^*,n}$ are positive,

$$\delta_n = \frac{\Delta'_{k,m} w_{n,m} / \tau_{k,n}}{\sum_{n' \in [M]} w_{n',m}^2 \left(\frac{1}{\tau_{k,n'}} + \frac{1}{\tau_{k_m^*,n'}} \right)} \text{ and } \delta'_n = \frac{\Delta'_{k,m} w_{n,m} / \tau_{k_m^*,n}}{\sum_{n' \in [M]} w_{n',m}^2 \left(\frac{1}{\tau_{k,n'}} + \frac{1}{\tau_{k_m^*,n'}} \right)}.$$

The conclusion follows by plugging these expressions to get the expression of the infimum.

A.2 Regret Lower Bound: Proof of Theorem 3

Given a bandit instance μ , we can consider two sets of possible changes of distributions: changes that are allowed to change the distributions of optimal arms $(k_m^*)_{m \in [M]}$, and those that cannot

$$\begin{aligned} \text{Alt}(\mu) &:= \left\{ \lambda = (\lambda_{k,m})_{k,m} \in \mathbb{R}^{K \times M} : \exists m \in [M], k \neq k_m^* : \sum_n \lambda_{k,m} w_{n,m} > \sum_n \lambda_{k_m^*,n} w_{n,m} \right\}, \\ B(\mu) &:= \text{Alt}(\mu) \cap \left\{ \lambda = (\lambda_{k,m})_{k,m} \in \mathbb{R}^{K \times M} : \forall m \in [M], \lambda_{k_m^*,m} = \mu_{k_m^*,m} \right\}. \end{aligned}$$

For cumulative regret, the change-of-distribution lemma becomes an asymptotic result, stated below

Lemma 6. Fix $\mu \in \mathbb{R}^{K \times M}$, and let us consider $\lambda \in \text{Alt}(\mu)$. Then, for any $\varepsilon > 0$, there exists $T_0 = T_0(\mu, \lambda, \varepsilon)$, such that, for any $T \geq T_0$

$$\sum_{m,k} \mathbb{E}_\mu [N_{k,m}(T)] \frac{(\mu_{k,m} - \lambda_{k,m})^2}{2} \geq (1 - \varepsilon) \log(T).$$

Proof. The proof uses a change-of-distribution, following a technique proposed by [16]. Using the data processing inequality, and letting \mathcal{H}_T be the observations available to the central server (which sees all local rewards under our assumptions), we have that

$$\text{KL}(\mathbb{P}_\mu^{\mathcal{H}_T}, \mathbb{P}_\lambda^{\mathcal{H}_T}) \geq \text{kl}(\mathbb{P}_\mu(\mathcal{E}_T), \mathbb{P}_\lambda(\mathcal{E}_T)),$$

where KL is the Kullback-Leibler divergence and $\mathbb{P}_\mu^{\mathcal{H}_T}$ is the distribution of the observation under the bandit model μ , and \mathcal{E}_T is any event. Using that

$$\text{KL}(\mathbb{P}_\mu^{\mathcal{H}_T}, \mathbb{P}_\lambda^{\mathcal{H}_T}) = \sum_{k,m} \mathbb{E}_\mu [N_{k,m}(T)] \frac{(\mu_{k,m} - \lambda_{k,m})^2}{2}$$

together with the lower bound $\text{kl}(p, q) \geq (1 - p) \log(1/(1 - q)) - \log(2)$, where kl is the binary relative entropy and for any distributions p, q , yields

$$\sum_{k,m} \mathbb{E}_\mu [N_{k,m}(T)] \frac{(\mu_{k,m} - \lambda_{k,m})^2}{2} \geq (1 - \mathbb{P}_\mu(\mathcal{E}_T)) \log\left(\frac{1}{\mathbb{P}_\lambda(\mathcal{E}_T)}\right) - \log(2).$$

We now pick the event

$$\mathcal{E}_T = \left\{ N_{k_m^*, T}(T) \leq \frac{T}{2} \right\},$$

and use that \mathcal{E}_T is very unlikely under μ as for any $\gamma \in (0, 1)$,

$$\mathbb{P}_\mu(\mathcal{E}_T) = \mathbb{P}_\mu \left(\sum_{k \neq k_m^*} N_{k,m}(T) \geq \frac{T}{2} \right) = \frac{2 \sum_{k \neq k_m^*} \mathbb{E}_\mu [N_{k,m}(T)]}{T} = \frac{o_{T \rightarrow \infty}(T^\gamma)}{T},$$

and very likely under any λ for which k_m^* is suboptimal as, for any $\gamma \in (0, 1)$,

$$\mathbb{P}_\lambda(\overline{\mathcal{E}_T}) = \mathbb{P}_\lambda \left(N_{k_m^*, T}(T) > \frac{T}{2} \right) = \frac{2 \mathbb{E}_\lambda [N_{k_m^*, m}(T)]}{T} = \frac{o_{T \rightarrow \infty}(T^\gamma)}{T},$$

where we exploit the fact that the algorithm is uniformly efficient (its regret and therefore its number of sub-optimal draws is $o(T^\gamma)$ under any bandit model). The conclusion follows from some elementary real analysis to prove that the right hand side of the inequality is larger than $(1 - \varepsilon) \log(T)$ for T large enough (how larger depends in a complex way of $\mu, \lambda, \varepsilon$ and the algorithm). \square

At this point, we would really like to select the alternative model λ that leads to the tightest inequality in Lemma 6. For example, using Lemma 7 below, we find that the best way to make an arm $k \neq k_m^*$ better than k_m^* consists in choosing

$$\lambda_{k,n}(T) = \mu_{k,n} + \frac{\Delta_{k,m} w_{n,m} / \mathbb{E}_\mu [N_{k,m}(T)]}{\sum_{m' \in [M]} \frac{w_{n',m}^2}{\mathbb{E}_\mu [N_{k,n'}(T)]}},$$

for any $n \in [M]$ and $\lambda_{k',n}(T) = \mu_{k',n}$ for any arm $k' \neq k$. However, this choice of alternative λ depends on T , hence we cannot apply Lemma 6, which is asymptotic in T and holds for a fixed λ . We have to be careful to be able to exchange the \liminf over T and the infimum over alternatives in the constraints, and we will be able to do so only for changes of measures that are restricted to $B(\mu)$.

We first assume that $\liminf_{T \rightarrow \infty} \frac{\mathcal{R}(T)}{\log(T)}$ is finite, and call its value $\ell(\mu)$ (this is fine as otherwise any lower bound trivially holds). By definition of the \liminf , there exists a sequence $(T_i)_{i \in \mathbb{N}}$ such that

$$\liminf_{T \rightarrow \infty} \frac{\mathcal{R}(T)}{\log(T)} = \lim_{i \rightarrow \infty} \sum_{m \in [M], k \neq k_m^*} \Delta'_{k,m} \frac{\mathbb{E}_\mu[N_{k,m}(T_i)]}{\log(T_i)} = \ell(\mu).$$

The fact that this sequence has a limit, and that the gaps are positive, implies that each sequence $(\mathbb{E}_\mu[N_{k,m}(T_i)]/\log(T_i))_{i \in \mathbb{N}}$ is bounded. Therefore, there must exist a subsequence, that we denote $(T'_i)_{i \in \mathbb{N}}$ of $(T_i)_{i \in \mathbb{N}}$ such that

$$\forall m \in [M], k \neq k_m^*, \lim_{i \rightarrow \infty} \frac{\mathbb{E}_\mu[N_{k,m}(T'_i)]}{\log(T'_i)} = c_{k,m},$$

for some real values $(c_{k,m})_{k \in [K], m \in [M]}$, and, in particular,

$$\ell(\mu) = \sum_{m \in [M], k \neq k_m^*} \Delta'_{k,m} c_{k,m}.$$

Now, it follows from Lemma 6 that, for any $\lambda \in \text{Alt}(\mu)$,

$$\liminf_{T \rightarrow \infty} \sum_{k,m} \frac{\mathbb{E}_\mu[N_{k,m}(T)]}{\log(T)} \frac{(\mu_{k,m} - \lambda_{k,m})^2}{2} \geq 1.$$

But we have no idea about the behavior of the sequence $(\mathbb{E}_\mu[N_{k,m}(T)]/\log(T))_{T \in \mathbb{N}}$ for $k = k_m^*$, this is why we have to consider only $\lambda \in B(\mu)$, for which we deduce that

$$\begin{aligned} \lim_{i \rightarrow \infty} \sum_{m \in [M], k \neq k_m^*} \frac{\mathbb{E}_\mu[N_{k,m}(T'_i)]}{\log(T'_i)} \frac{(\mu_{k,m} - \lambda_{k,m})^2}{2} &\geq 1, \\ \sum_{m \in [M], k \neq k_m^*} c_{k,m} \frac{(\mu_{k,m} - \lambda_{k,m})^2}{2} &\geq 1. \end{aligned}$$

(the \liminf being the lowest value of the limit of any subsequence). This proves that

$$\ell(\mu) \geq \min_c \sum_{m \in [M], k \neq k_m^*} \Delta'_{k,m} c_{k,m},$$

under the constraints that $c = (c_{k,m})_{k,m:k \neq k_m^*}$ belongs to the constraint set

$$\mathcal{C} = \left\{ (c_{k,m})_{k,m:k \neq k_m^*} : \forall \lambda \in B(\mu), \sum_{m \in [M], k \neq k_m^*} c_{k,m} \frac{(\mu_{k,m} - \lambda_{k,m})^2}{2} \geq 1 \right\}.$$

We now establish that

$$\mathcal{C} \subseteq \bigcup_{k,m} \left\{ (c_{k,n})_{k,n:k \neq k_n^*} : \sum_{n:k \neq k_n^*} \frac{w_{n,m}^2}{c_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2} \right\}, \quad (5)$$

by selecting some well chosen elements in $B(\mu)$.

First, for every $(m, k) \in [M] \times [K]$ such that $k \neq k_m^*$, for every $\delta = (\delta_n)_{n:k_n^* \neq k}$, we define an instance λ^δ by $\lambda_{k,n}^\delta = \mu_{k,n} + \delta_n$ for any $n \in [M]$ such that $k \neq k_n^*$, and $\lambda_{k,n}^\delta = \mu_{k,n}$ otherwise. We observe that λ^δ belongs to $B(\mu)$ if

$$\sum_{n \in [M]: k \neq k_n^*} w_{n,m} \delta_n > \Delta'_{k,m},$$

as this leads to $\sum_{n \in [M]} w_{n,m} \lambda_{k,n} > \sum_{n \in [M]} w_{n,m} \lambda_{k_m^*,n}$, and arm k_m^* being sub-optimal in λ^δ . For all $c \in \mathcal{C}$,

$$\min_{\delta: \sum_{n:k \neq k_n^*} w_{n,m} \delta_n \geq \Delta'_{k,m}} \sum_{n \in [M]: k_n^* \neq k} c_{k,n} \frac{\delta_n^2}{2} \geq 1.$$

From Lemma 7, this leads to

$$\sum_{n \in [M]: k_n^* \neq k} \frac{w_{n,m}^2}{c_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2}.$$

Now we consider $(m, k) \in [M] \times [K]$, such that $k = k_m^*$. In this case, we define an instance λ^δ by $\lambda_{k_m^*, n}^\delta = \mu_{k_m^*, n} - \delta_n$ for any $n \in [M]$ such that $k_m^* \neq k_n^*$, and $\lambda_{k,n}^\delta = \mu_{k,n}$ otherwise. We observe that λ^δ belongs to $B(\mu)$ if

$$\sum_{n \in [M]: k \neq k_n^*} w_{n,m} \delta_n > \min_{k' \in [K]: k' \neq k_m^*} \Delta'_{k',m} = \Delta'_{k_m^*,m},$$

as this leads to $\sum_{n \in [M]} w_{n,m} \lambda_{k_m^*, n} < \max_{k' \neq k_m^*} \sum_{n \in [M]} w_{n,m} \mu_{k', n}$ and arm k_m^* being sub-optimal in λ^δ . For any $c \in \mathcal{C}$,

$$\min_{\delta: \sum_{n: k \neq k_n^*} w_{n,m} \delta_n \geq \Delta'_{k_m^*,m}} \sum_{n: k_n^* \neq k_m^*} c_{k,n} \frac{\delta_n^2}{2} \geq 1,$$

and Lemma 7 leads to

$$\sum_{n \in [M]: k_n^* \neq k_m^*} \frac{w_{n,m}^2}{c_{k,n}} \leq \frac{(\Delta'_{k_m^*,m})^2}{2}.$$

In conclusion, for $c \in \mathcal{C}$, we proved that, for every $(k, m) \in [K] \times [M]$,

$$\sum_{n \in [M]: k_n^* \neq k} \frac{w_{n,m}^2}{c_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2},$$

which proves the inclusion (5) and concludes the proof, as the minimum over a larger set is (potentially) smaller.

Lemma 7. *Let \mathcal{N} be a set of indices, and $(c_n)_{n \in \mathcal{N}}$ be all positive. The minimizer over $\delta \in \mathbb{R}^{|\mathcal{N}|}$ of $\sum_{n \in \mathcal{N}} c_n \frac{\delta_n^2}{2}$, under the constraint that $\sum_{n \in \mathcal{N}} \delta_n w_{n,m} \geq d_m$, satisfies*

$$\forall n \in \mathcal{N}, \delta_n = \frac{d_m w_{n,m} / c_n}{\sum_{n' \in \mathcal{N}} \frac{w_{n',m}^2}{c_{n'}}},$$

and the minimum is equal to

$$\frac{d_m^2}{2} \left(\sum_{n \in \mathcal{N}} \frac{w_{n,m}^2}{c_n} \right)^{-1}.$$

Lemma 7 is proven using classical techniques for solving constrained minimization problems.

B Proof of Theorem 2

We recall the good event \mathcal{E} defined in Lemma 3. First, the following lemma ensures the correctness of Algorithm 1 on \mathcal{E} , which holds with probability $1 - \delta$ from Lemma 3.

Lemma 8. *On event \mathcal{E} , when stopping, W-CPE-BAI outputs $\hat{k}_m = k_m^*$ for each agent m .*

Proof. On event \mathcal{E} , it is not possible that one agent m eliminates arm k_m^* from its set $B_m(r+1)$ at any round r ; otherwise, if $j_m(r) \in \arg \max_{j \in B_m(r)} \{\hat{\mu}'_{j,m}(r) - \Omega_{j,m}(r)\}$, $j_m(r) \neq k_m^*$, then the elimination criterion and event \mathcal{E} imply that $\mu'_{k_m^*,m} < \mu'_{j_m(r),m}$, which is absurd. \square

Then we upper bound the exploration cost when \mathcal{E} holds. We denote by R the (random) number of rounds used by the algorithm, and, for all $m \in [M]$ and $k \neq k_m^*$, by $R_{k,m}$ the (random) last round in which k is still a candidate arm for player m

$$R_{k,m} := \sup \{r \geq 0 : k \in B_m(r)\}.$$

By definition of Algorithm 1, $R = \max_m \max_{k \neq k_m^*} R_{k,m}$. We first provide upper bounds on $R_{k,m}$ and R . To achieve this, we introduce the following notation for any $k \in [K], m \in [M]$,

$$r_{k,m} := \min \{r \geq 0 : 4 \times 2^{-r} < \Delta'_{k,m}\} \text{ and } r_{\max} := \max_{m \in [M]} \max_{k \neq k_m^*} r_{k,m}.$$

The following upper bounds can be easily checked.

Lemma 9. $\forall k, m, r_{k,m} \leq \log_2(8/\Delta'_{k,m})$ and $r_{\max} \leq \log_2(8/\Delta'_{\min})$.

Using the fact that W-CPE -BAI is only halving the proxy gaps of arms that are not eliminated, we can write down the value of the proxy gaps for these arms.

Lemma 10. $\forall m \in [M]$ and $k \in B_m(r)$, $\tilde{\Delta}_{k,m}(r) = 2^{-r}$.

Using the important relationship between proxy gaps and the confidence width as established in Lemma 4, we can further show that

Lemma 11. On \mathcal{E} , for any $m \in [M], k \neq k_m^*, R_{k,m} \leq r_{k,m}$.

Proof. Assume \mathcal{E} holds. For any suboptimal arm k for agent m , at round $r = r_{k,m}$, if $k \notin B_m(r)$, then $R_{k,m} < r_{k,m}$. Otherwise, if $k \in B_m(r)$, then we know that $k_m^* \in B_m(r)$, as event \mathcal{E} holds (see the proof of Lemma 8). Then

$$\begin{aligned} \hat{\mu}'_{k,m}(r) + \Omega_{k,m}(r) &\leq_{(1)} \mu'_{k,m} + 2\Omega_{k,m}(r) \\ &\leq_{(2)} \mu'_{k,m} + 2\tilde{\Delta}_{k,m}(r) = \mu'_{k,m} + 4\tilde{\Delta}_{k,m}(r) - 2\tilde{\Delta}_{k,m}(r) \\ &<_{(3)} \mu'_{k,m} + \Delta'_{k,m} - 2\tilde{\Delta}_{k,m}(r) = \mu'_{k_m^*,m} - 2\tilde{\Delta}_{k,m}(r) \\ &\leq_{(1)} \hat{\mu}'_{k_m^*,m}(r) + \Omega_{k_m^*,m}(r) - 2\tilde{\Delta}_{k,m}(r) \\ &\leq_{(2),(4)} \max_{j \in B_m(r)} \{\hat{\mu}'_{j,m}(r) - \Omega_{j,m}(r)\} + 2 \cdot 2^{-r} - 2 \cdot 2^{-r}, \\ \implies \hat{\mu}'_{k,m}(r) + \Omega_{k,m}(r) &< \max_{j \in B_m(r)} \{\hat{\mu}'_{j,m}(r) - \Omega_{j,m}(r)\}, \end{aligned}$$

where (1) is using that event \mathcal{E} holds ; (2) is using Lemma 4 ; (3) is using the definition of $r = r_{k,m}$, the fact that $k \in B_m(r)$ and Lemma 10 and (4) is using that $k, k_m^* \in B_m(r)$ and Lemma 10. It follows that $k \notin B_m(r_{k,m} + 1)$ and $R_{k,m} \leq r_{k,m}$. \square

The previous lemma straightforwardly implies that

Corollary 1. $R \leq r_{\max} \leq \log_2(8/\Delta'_{\min})$.

Moreover, it also permits to prove that, in the last round R , the proxy gaps are lower bounded by the gaps.

Corollary 2. At final round R , and for any agent m and suboptimal (with respect to m) arm $k \neq k_m^*$, if \mathcal{E} holds,

$$\tilde{\Delta}_{k,m}(R) \geq \frac{1}{8} \Delta'_{k,m}.$$

Proof. If $R < r_{k,m}$, by definition of $r_{k,m}$, we have that $\tilde{\Delta}_{k,m}(R) \geq (1/4)\Delta'_{k,m} \geq (1/8)\Delta'_{k,m}$. If $R \geq r_{k,m}$, we first observe that $\tilde{\Delta}_{k,m}(R) = \tilde{\Delta}_{k,m}(R_{k,m}) = (1/2)\tilde{\Delta}_{k,m}(R_{k,m} - 1)$ by definition of the algorithm (the gaps remain frozen when an arm is eliminated, and they are halved otherwise). As $R_{k,m} - 1 < r_{k,m}$ by Lemma 11, by definition of $r_{k,m}$, it follows that

$$4\tilde{\Delta}_{k,m}(R_{k,m} - 1) > \Delta'_{k,m}$$

and we conclude that $\tilde{\Delta}_{k,m}(R) \geq (1/8)\Delta'_{k,m}$. \square

Note that, for any $m \in [M]$, $\tilde{\Delta}_{k^*,m}(R) = \min_{k \neq k^*} \tilde{\Delta}_{k,m}(R)$ by Algorithm 1. Now, for any $m \in [M], k \in [K]$, using Corollary 2, and the fact that the proxy gaps are non-increasing between two consecutive phases, we get

$$\forall k \in [K], \forall m \in [M], \forall r \leq R, \tilde{\Delta}_{k,m}(r) \geq \tilde{\Delta}_{k,m}(R) \geq \frac{\Delta'_{k,m}}{8}.$$

Using Lemma 2 and the fact that proxy gaps are non-increasing, for any round $r \leq R$, the optimal allocation $t(r) \in \tilde{\mathcal{P}}^* \left(\left(\sqrt{2} \tilde{\Delta}_{k,m}(r) \right)_{k,m} \right)$ satisfies

$$\sum_{k,m} t_{k,m}(r) \leq 32 \sum_{k,m} t'_{k,m},$$

where $t' \in \tilde{\mathcal{P}}^*(\Delta')$, hence

$$\max_{r \leq R} \left[\sum_{k,m} t_{k,m}(r) \right] \leq 32 \tilde{T}_W^*(\mu). \quad (6)$$

For every $k \in [K], m \in [M]$, we now introduce

$$r'_{k,m} := \sup\{r \leq R : d_{k,m}(r) \neq 0\},$$

so that $n_{k,m}(R) = n_{k,m}(r'_{k,m})$. Using Lemma 12 stated below, and the fact that function β_δ is nondecreasing in each coefficient of its argument (see its definition in Lemma 3),

$$\begin{aligned} n_{k,m}(R) = n_{k,m}(r'_{k,m}) &\leq t_{k,m}(r'_{k,m}) \beta_\delta(n_{k,\cdot}(r'_{k,m})) + 1 \\ &\leq t_{k,m}(r'_{k,m}) \beta_\delta(n_{k,\cdot}(R)) + 1. \end{aligned}$$

Lemma 12. For any $k, m, r \geq 0$, either $d_{k,m}(r) = 0$, or $n_{k,m}(r) = n_{k,m}(r-1) + d_{k,m}(r) < t_{k,m}(r) \beta_\delta(n_{k,\cdot}(r)) + 1$.

Proof. At fixed $r \geq 0$, for any set $S \subseteq [K] \times [M]$, let us prove by induction on $|S| \geq 1$ ⁷

$$\begin{aligned} \forall k, m, \quad & d'_{k,m}(r) := (d_{k,m}(r) - \mathbb{1}_S((k, m)))_+ \\ \implies \quad & \forall (k, m) \in S, \frac{n_{k,m}(r-1) + d'_{k,m}(r)}{\beta_\delta(n_{k,\cdot}(r-1) + d_{k,\cdot}(r))} < t_{k,m}(r) \\ & \text{or } d_{k,m}(r) = 0. \end{aligned}$$

At $|S| = 1$: Let us denote $S = \{(k', m')\}$. If $d_{k',m'}(r) = 0$, then it is trivial. Otherwise, $\sum_{k,m} d'_{k,m}(r) < \sum_{k,m} d_{k,m}(r)$, and then, by minimality of solution $d(r)$, at least one constraint from the optimization problem of value $\sum_{k,m} d_{k,m}(r)$ has to be violated. For any $(k, m) \notin S$, by definition of $d(r)$ and nondecreasingness of β_δ ,

$$\begin{aligned} n_{k,m}(r-1) + d'_{k,m}(r) &= n_{k,m}(r-1) + d_{k,m}(r) \\ &\geq t_{k,m}(r) \beta_\delta(n_{k,\cdot}(r-1) + d_{k,\cdot}(r)) \\ &\geq t_{k,m}(r) \beta_\delta(n_{k,\cdot}(r-1) + d'_{k,\cdot}(r)). \end{aligned}$$

That means, necessarily the only constraint that is violated is the one on (k', m') . Using the nondecreasingness of β_δ :

$$\begin{aligned} n_{k',m'}(r-1) + d_{k',m'}(r) - 1 &= n_{k',m'}(r-1) + d'_{k',m'}(r) \\ &< t_{k',m'}(r) \beta_\delta(n_{k',\cdot}(r-1) + d'_{k',\cdot}(r)) \\ &\leq t_{k',m'}(r) \beta_\delta(n_{k',\cdot}(r-1) + d_{k',\cdot}(r)). \end{aligned}$$

Combining the two ends of the inequality proves the claim.

⁷For any $x \in \mathbb{N}^M$, $(x)_+ := (\max(0, x_m))_{m \in [M]}$, and $\mathbb{1}_S$ is the indicator function of set S .

At $|S| > 1$: At fixed $(k', m') \in S$, we can apply the claim to $S \setminus \{(k', m')\}$. Moreover, if $d_{k', m'}(r) = 0$, then the claim is proven. Otherwise, by appealing to the extremes,

$$n_{k', m'}(r-1) + d'_{k', m'}(r) \geq t_{k', m'}(r) \beta_\delta(n_{k', \cdot}(r-1) + d_{k, \cdot}(r)).$$

Let us then consider the following allocation:

$$\forall k, m, d''_{k, m}(r) := d_{k, m}(r) - \mathbb{1}_{\{(k', m')\}}((k, m)).$$

It can be checked straightforwardly – using the nondecreasingness of β_δ – that d'' satisfies all required constraints for any pair $(k, m) \in [K] \times [M]$, and that $\sum_{k, m} d''_{k, m}(r) = \sum_{k, m} d_{k, m}(r) - 1$, which, by minimality of d , is absurd. Then the claim is proven for $|S| > 1$. Then Lemma 12 is proven by considering $S = [K] \times [M]$. \square

By summing the upper bound on $(n_{k, m}(R))_{k, m}$ over $[K] \times [M]$, we can upper bound the exploration cost τ as

$$\begin{aligned} \tau := \sum_{k, m} n_{k, m}(R) &\leq \sum_{k, m} t_{k, m}(r'_{k, m}) \beta_\delta(n_{k, \cdot}(R)) + KM \\ &\leq \sum_{k, m} t_{k, m}(r'_{k, m}) \beta^*(\tau) + KM \\ &\leq \sum_{k, m} \sum_{r \leq R} t_{k, m}(r) \beta^*(\tau) + KM \\ &\leq R \max_{r \leq R} \left[\sum_{k, m} t_{k, m}(r) \right] \beta^*(\tau) + KM \\ &\leq \log_2(8/\Delta'_{\min}) \max_{r \leq R} \left[\sum_{k, m} t_{k, m}(r) \right] \beta^*(\tau) + KM \end{aligned}$$

where we use Corollary 1 and introduce the quantity

$$\beta^*(\tau) := \beta_\delta(\tau \mathbb{1}_M) = 2 \left(g_M \left(\frac{\delta}{KM} \right) + 2M \ln(4 + \ln(\tau)) \right) \text{ where } \forall n \in [M], \mathbb{1}_M(n) = 1.$$

Using Equation (6) and Lemma 1,

$$\tau \leq 32 \tilde{T}_W^*(\mu) \log_2(8/\Delta'_{\min}) \beta^*(\tau) + KM \leq 32 T_W^*(\mu) \log_2(8/\Delta'_{\min}) \beta^*(\mu) + KM.$$

Therefore, τ is upper bounded by

$$\sup \{ n \in \mathbb{N}^* : n \leq 32 T_W^*(\mu) \log_2(8/\Delta'_{\min}) \beta^*(n) + KM \}.$$

Applying Lemma 15 in [24] with

$$\begin{aligned} \Delta &= \left(\sqrt{32 T_W^*(\mu) \log_2(8/\Delta'_{\min})} \right)^{-1}, \\ a &= KM + 2g_M \left(\frac{\delta}{KM} \right), \\ b &= 4M, \\ c &= 4, \\ d &= e^{-1} \text{ (using } \forall n, \log(n) \leq ne^{-1} \text{)}, \end{aligned}$$

τ is upper bounded by

$$\begin{aligned} \hat{T}_W(\mu) &:= 32 T_W^*(\mu) \log_2(8/\Delta'_{\min}) \left[KM + 2g_M \left(\frac{\delta}{KM} \right) \right. \\ &\quad \left. + 4M \ln \left(4 + 1,024 \frac{(T_W^*(\mu) \log_2(8/\Delta'_{\min}))^2}{e} \left(KM + 2g_M \left(\frac{\delta}{KM} \right) + 4M(2 + \sqrt{e}) \right)^2 \right) \right], \end{aligned}$$

which satisfies $\tau \hat{T}_W(\mu) \leq a + b \ln(c + d\tau \hat{T}_W(\mu))$. Using that $g_M(x) \simeq x + M \log \log(x)$ in the regime of small values of δ , we obtain that

$$\hat{T}_W(\mu) = 32 T_W^*(\mu) \log_2(8/\Delta'_{\min}) \log(1/\delta) + o_{\delta \rightarrow 0}(\log(1/\delta)).$$

The upper bound on the communication cost follows from the upper bound on the number of phases given in Corollary 1.

C Supplementary Lemmas and Proofs

We report here technical lemmas and their proofs, ordered by section.

C.1 Collaborative Best-Arm Identification

Lemma 13. *Introducing the quantity*

$$\tilde{T}_W^*(\mu) := \min_{t \in (\mathbb{R}^+)^{K \times M}} \sum_{(k,m) \in [K] \times [M]} t_{k,m} \text{ s.t. } \forall m \in [M], k \in [K], \sum_{n \in [M]} \frac{w_{n,m}^2}{t_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2},$$

it holds that $\tilde{T}_W^*(\mu) \leq T_W^*(\mu) \leq 2\tilde{T}_W^*(\mu)$.

Proof. Let us denote by \mathcal{C} and $\tilde{\mathcal{C}}$ the two constraint sets such that $T_W^*(\mu) = \min \{ \sum_{k,m} t_{k,m} \mid t \in \mathcal{C} \}$ and $\tilde{T}_W^*(\mu) = \min \{ \sum_{k,m} t_{k,m} \mid t \in \tilde{\mathcal{C}} \}$. The inequality $\tilde{T}_W^*(\mu) \leq T_W^*(\mu)$ is obtained by noticing that $\mathcal{C} \subseteq \tilde{\mathcal{C}}$. To prove the other inequality, we consider $\tilde{\tau} \in \arg \min \{ \sum_{k,m} t_{k,m} \mid t \in \tilde{\mathcal{C}} \}$. Then, for any agent $m \in [M]$, arm $k \neq k_m^*$,

$$\begin{aligned} \sum_{n \in [M]} w_{n,m}^2 \left(\frac{1}{2\tilde{\tau}_{k,n}} + \frac{1}{2\tilde{\tau}_{k_m^*,n}} \right) &= \frac{1}{2} \left(\sum_{n \in [M]} \frac{w_{n,m}^2}{\tilde{\tau}_{k,n}} \right) + \frac{1}{2} \left(\sum_{n \in [M]} \frac{w_{n,m}^2}{\tilde{\tau}_{k_m^*,n}} \right) \\ &\leq \underbrace{(\Delta'_{k,m})^2/2}_{\leq (\Delta'_{k_m^*,m})^2/2} + \underbrace{(\Delta'_{k_m^*,m})^2/2}_{\leq (\Delta'_{k',m})^2/2} = \min \{ (\Delta'_{k',m})^2/2 \mid k' \neq k_m^* \} \\ &\leq (\Delta'_{k,m})^2/2. \end{aligned}$$

Then $2\tilde{\tau} \in \mathcal{C}$, therefore by minimality, $T_W^*(\mu) \leq 2\tilde{T}_W^*(\mu)$. \square

Lemma 14. *Consider $\Delta, \Delta' \in (\mathbb{R}^+)^{K \times M}$, such that $\tau \in \tilde{\mathcal{P}}^*(\Delta)$ and $\tau' \in \tilde{\mathcal{P}}^*(\Delta')$. Then*

(i). *If there exists $\alpha > 0$ such that: $\forall k \in [K], \forall m \in [M], \alpha \Delta_{k,m} \leq \Delta'_{k,m}$,*

$$\sum_{k,m} \tau'_{k,m} \leq \frac{1}{\alpha^2} \sum_{k,m} \tau_{k,m}.$$

(ii). *If there exists $\beta > 0$ such that: $\forall k \in [K], \forall m \in [M], \Delta'_{k,m} \leq \beta \Delta_{k,m}$. Then*

$$\frac{1}{\beta^2} \sum_{k,m} \tau_{k,m} \leq \sum_{k,m} \tau'_{k,m}.$$

Proof. The proof follows from the fact that τ and τ' are *minimal*. In particular, to prove (ii), let $\tau''_{k,m} = \beta^2 \tau'_{k,m}$ for any $k \in [K], m \in [M]$. Then, for any agent m and arm k ,

$$\sum_{n \in [M]} \frac{w_{n,m}^2}{\tau''_{k,n}} = \sum_{n \in [M]} \frac{w_{n,m}^2}{\beta^2 \tau'_{k,n}} \leq \frac{1}{2} \left(\frac{\Delta'_{k,m}}{\beta} \right)^2 \leq \frac{\Delta_{k,m}^2}{2}.$$

By minimality of τ ,

$$\sum_{m \in [M]} \tau_{k,m} \leq \sum_{m \in [M]} \tau''_{k,m} = \beta^2 \sum_{m \in [M]} \tau'_{k,m}.$$

(i) similarly follows. \square

C.2 A Near-Optimal Algorithm For Best Arm Identification

Lemma 15. *Let us define*

$$\beta_\delta(N) := 2 \left(g_M \left(\frac{\delta}{KM} \right) + 2 \sum_{m=1}^M \ln(4 + \ln(N_m)) \right),$$

for any $N \in (\mathbb{N}^*)^M$, where

$$\forall \delta \in (0, 1), g_M(\delta) := M \mathcal{C}^{g_G}(\log(1/\delta)/M),$$

$$\forall x > 0, \mathcal{C}^{g_G}(x) := \min_{\lambda \in (0.5, 1)} \frac{g_G(\lambda) + x}{\lambda},$$

$$\text{and } \forall \lambda \in (0.5, 1), g_G(\lambda) := 2\lambda - 2\lambda \log(4\lambda) + \log(\zeta(2\lambda)) - 0.5 \log(1 - \lambda),$$

where ζ is the Riemann zeta function. Then, the good event

$$\mathcal{E} := \left\{ \forall r \in \mathbb{N}, \forall m, \forall k, |\hat{\mu}'_{k,m}(r) - \mu'_{k,m}| \leq \Omega_{k,m}(r) \right\}.$$

holds with probability larger than $1 - \delta$.

Proof. Using Proposition 24 from [23] on $\mu'_{k,m}$, for any arm k and agent m , directly yields

$$\mathbb{P} \left(\exists r \geq 0, |\hat{\mu}'_{k,m}(r) - \mu'_{k,m}| > \sqrt{2 \left(g_M \left(\frac{\delta}{KM} \right) + 2 \sum_{n=1}^M \ln(4 + \ln(n_{k,n}(r))) \right) \sum_n \frac{w_{n,m}^2}{n_{k,n}(r)}} \right) \leq \frac{\delta}{KM}$$

(using the notation of the paper, consider $\mu = \mu_{k,\cdot}$ and $c = W_{\cdot,m}$). Then all that is needed to conclude is to apply a union bound on $[K] \times [M]$

$$\begin{aligned} \mathbb{P}(\mathcal{E}^c) &= \mathbb{P} \left(\exists m \in [M], \exists k \in [K], \exists r \geq 0, |\hat{\mu}'_{k,m}(r) - \mu'_{k,m}| > \sqrt{2\beta_\delta(n_{k,\cdot}(r)) \sum_n \frac{w_{n,m}^2}{n_{k,n}(r)}} \right) \\ &\leq \sum_{m \in [M]} \sum_{k \in [K]} \frac{\delta}{KM} \leq \delta. \end{aligned}$$

□

C.3 Regret Lower Bound

Lemma 16. *Introducing the quantity*

$$\tilde{C}_W^*(\mu) := \min_{c \in (\mathbb{R}^+)^{K \times M}} \sum_{k=1}^K \sum_{m=1}^M c_{k,m} \Delta'_{k,m} \text{ s.t. } \forall k \in [K], \forall m \in [M], \sum_{n=1}^M \frac{w_{n,m}^2}{c_{k,n}} \leq \frac{\Delta'_{k,m}}{2\sigma^2},$$

it holds that $C_W^*(\mu) \leq \tilde{C}_W^*(\mu) \leq 4C_W^*(\mu)$.

Proof. Let c and \tilde{c} be the solutions to the optimization problems of $C_W^*(\mu)$ and $\tilde{C}_W^*(\mu)$, respectively. Note that, for any agent m , $c_{m,k_m^*} = +\infty$ because, in the optimization problem related to the regret lower bound, these terms do not contribute to the objective. The lower bound follows from the definition of c and the fact that $(\tilde{c}_{k,m})_{m,k \neq k_m^*}$ satisfy the same constraints as $(c_{k,m})_{m,k \neq k_m^*}$.

Next, we prove the upper bound on $\tilde{C}_W^*(\mu)$. For any $k \in [K]$, define $\mathcal{S}_k^* = \{m : k = k_m^*\}$ and $\mathcal{S}_k = \{m : k \neq k_m^*\}$ (note that for any $k \in [K]$, $\{\mathcal{S}_k^*, \mathcal{S}_k\}$ is a partition of $[M]$). For $k \in [K]$ and $m \in \mathcal{S}_k$, let $c'_{k,m} = 2c_{k,m}$. For $k \in [K]$ and $m \in \mathcal{S}_k^*$, let $c'_{k,m} = c_{k,m}^1$, where

$$c^1 \in \arg \min_{\tau \in (\mathbb{R}^+)^{K \times M}} \sum_{k=1}^K \sum_{m \in \mathcal{S}_k^*} \tau_{k,m} \Delta'_{k,m} \text{ s.t. } \forall k \in [K], m \in [M], \sum_{n \in \mathcal{S}_k^*} \frac{w_{n,m}^2}{\tau_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{4}.$$

Note that $c'_{k,m}$ satisfy the same constraints as $\tilde{c}_{k,m}$: $\forall k \in [K], m \in [M], \sum_{n=1}^M \frac{w_{n,m}^2}{c'_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2}$. We thus have

$$\begin{aligned}
\sum_{k=1}^K \sum_{m=1}^M \tilde{c}_{k,m} \Delta'_{k,m} &\leq \sum_{k=1}^K \sum_{m=1}^M c'_{k,m} \Delta_{k,m} \\
&= \sum_{k=1}^K \sum_{m \in \mathcal{S}_k} c'_{k,m} \Delta_{k,m} + \sum_{k=1}^K \sum_{m \in \mathcal{S}_k^*} c'_{k,m} \Delta_{k,m} \\
&= 2 \sum_{k=1}^K \sum_{m \in \mathcal{S}_k} c_{k,m} \Delta'_{k,m} + \sum_{k=1}^K \sum_{m \in \mathcal{S}_k^*} c'_{k,m} \Delta_{k,m} \\
&= 2C_W^*(\mu) + \sum_{k=1}^K \sum_{m \in \mathcal{S}_k^*} c_{k,m}^1 \Delta'_{k,m} \\
&\leq 4C_W^*(\mu).
\end{aligned}$$

The first inequality holds by minimality of \tilde{c} . To prove the second inequality, for all $m \in [M]$, let us define $c''_{k_m^*, m} = 2 \sum_{k \neq k_m^*} c_{k,m}$, and $c''_{k,m} = \infty$ for $k \neq k_m^*$. Note that c'' satisfy the constraints in the definition of c^1 . Thus, by minimality of c^1 , and using the fact that, for any $m \in [M]$, $\Delta'_{k_m^*, m} := \min_{k \neq k_m^*} \Delta'_{k,m}$, we have

$$\begin{aligned}
\sum_{k=1}^K \sum_{m \in \mathcal{S}_k^*} c_{k,m}^1 \Delta'_{k,m} &\leq \sum_{k=1}^K \sum_{m \in \mathcal{S}_k^*} c''_{k,m} \Delta'_{k,m} \\
&= 2 \sum_{k=1}^K \sum_{k, m \in \mathcal{S}_k^*} \left(\sum_{j \neq k_m^*} c_{j,m} \right) \Delta'_{k,m} \\
&= 2 \sum_{k=1}^K \sum_{m \in \mathcal{S}_k} c_{k,m} \Delta'_{k_m^*, m} \\
&\leq 2 \sum_{k=1}^K \sum_{m \in \mathcal{S}_k} c_{k,m} \Delta'_{k,m} \\
&= 2C_W^*(\mu),
\end{aligned}$$

which completes the proof. \square

D Extension to Top- N Identification

A generalization of the best arm identification problem is Top- N identification, which is the problem of finding the N optimal arms (for mixed rewards) for each agent. For any model $\mu \in \mathbb{R}^{K \times M}$, weight matrix $W \in [0, 1]^{M \times M}$, such that $\mu' = \mu W$, any agent m , and positive integer $N \leq K$, let us define⁸

$$\mathcal{S}_m^* := \left\{ k \in [K] \mid \mu'_{k,m} \geq \max_{k' \in [K]} \mu'_{k',m} = \max_{k' \in [K]} \sum_{n \in [M]} w_{n,m} \mu_{k',n} \right\}.$$

In this case, an algorithm for Top- N identification returns at the end of the exploration phase a set of N arms denoted $\hat{\mathcal{S}}_m$ for agent m . δ -correctness is defined as follows

⁸We define operation \max^N such that $\max^N_{i \in S} f(i)$ is the N^{th} (without multiplicity) greatest value in set $\{f(i) : i \in S\}$ for any function $f : S \mapsto \mathbb{R}$.

$$\mathbb{P}_\mu (\forall m \in [M], \hat{S}_m \subseteq S_m^*) \geq 1 - \delta .$$

Note that there might be more than N arms in a given set $S_m^*, m \in [M]$. Similarly to best arm identification, we also assume here that the set of top- N arms is unique – that is, for any m , $|S_m^*| = N$.

Lower Bound for Top- N Identification For Top- N identification, one can prove, similarly to the proof of Theorem 1 – using Lemma 1 in [28] to define the set of alternative models – the following result, which is valid for Gaussian rewards with fixed variance $\sigma^2 = 1$, and any weight matrix W such that all diagonal coefficients are positive,

Theorem 4. *Let μ be a fixed matrix of means in $\mathbb{R}^{K \times M}$. For any $\delta \in (0, 1/2]$, let \mathcal{A} be a δ -correct algorithm under which each agent communicates each reward to the central server after it is observed, and let us denote for any $k \in [K], m \in [M]$, $\tau_{k,m} := \mathbb{E}_\mu^{\mathcal{A}} [N_{k,m}(\tau)]$, where τ is the stopping time. For any $m \in [M], k \notin S_m^*, l \in S_m^*$, denoting $\mu' = \mu W$, it holds that*

$$\sum_n w_{n,m}^2 \left(\frac{1}{\tau_{k,n}} + \frac{1}{\tau_{l,n}} \right) \leq \frac{(\mu'_{k,m} - \mu'_{l,m})^2}{2 \log(1/(2.4\delta))} ,$$

and therefore $\text{Exp}_\mu(\mathcal{A}) \geq N_W^*(\mu) \log\left(\frac{1}{2.4\delta}\right)$, where

$$N_W^*(\mu) := \min_{t \in \mathbb{R}^{K \times M}} \sum_{(k,m) \in [K] \times [M]} t_{k,m}$$

$$\text{s.t. } \forall m, k \notin S_m^*(\mu), l \in S_m^*(\mu), \sum_{n \in [M]} w_{n,m}^2 \left(\frac{1}{t_{k,n}} + \frac{1}{t_{l,n}} \right) \leq \frac{(\mu'_{k,m} - \mu'_{l,m})^2}{2} .$$

Proof. As mentioned, let us use Lemma 1 from [28] to define the set of alternative models to μ in Top- N identification. If, for any agent m , $S_m^*(\mu)$ is the set of its top- N arms (of size N) with respect to mixed rewards, then

$$\begin{aligned} \text{Alt}(\mu) &:= \{ \lambda \in \mathbb{R}^{K \times M} : \exists m, S_m^*(\mu) \not\subseteq S_m^*(\lambda) \} \\ &= \{ \lambda \in \mathbb{R}^{K \times M} : \exists m, \exists k \notin S_m^*(\mu), \exists l \in S_m^*(\mu) : \lambda'_{k,m} > \lambda'_{l,m} \} , \end{aligned}$$

where $\lambda'_{k,m} := \sum_{n \in [M]} w_{n,m} \lambda_{k,n}$ for any arm k and agent m . If we assume that stopping time τ is almost surely finite under μ for algorithm \mathcal{A} , then let event $\mathcal{E}_\mu := \{ \exists m : \hat{S}_m \not\subseteq S_m^*(\mu) \}$. Using the δ -correctness of algorithm \mathcal{A} , where $\delta \leq 1/2$, by Theorem 1 from [15],

$$\frac{1}{2} \sum_{k,m} \tau_{k,m} (\mu_{k,m} - \lambda_{k,m})^2 \geq \log\left(\frac{1}{2.4\delta}\right) . \quad (7)$$

Similarly to the best arm identification case, we can show that, since all diagonal coefficients of W are positive, for any $k \in [K], m \in [M]$, $\tau_{k,m} > 0$. Consider now a fixed agent m , and two arms $k \notin S_m^*(\mu), l \in S_m^*(\mu)$. We will build an alternative model λ , similar enough to μ , where only arms k and l are modified for all agents, such that $l \notin S_m^*(\lambda)$ and $k \in S_m^*(\lambda)$. The procedure is similar to what we did in best arm identification. Given two nonnegative sequences $(\delta_n)_{n \in [M]}$ and $(\delta'_n)_{n \in [M]}$, we define $\lambda = (\lambda_{k',n})_{k' \in [K]}$ such that

$$\begin{cases} \lambda_{k',n} &= \mu_{k',n} \text{ if } k' \notin \{k, l\} , \\ \lambda_{k,n} &= \mu_{k,n} + \delta_n , \\ \lambda_{l,n} &= \mu_{l,n} - \delta'_n , \end{cases}$$

and which satisfies

$$(\lambda'_{k,m} - \mu'_{k,m}) - (\lambda'_{l,m} - \mu'_{l,m}) = \sum_{n \in [M]} w_{n,m} (\delta_n + \delta'_n) \geq \mu'_{l,m} - \mu'_{k,m} . \quad (8)$$

From Equation (7),

$$\inf_{\delta, \delta' : (8) \text{ holds}} \left[\sum_n \tau_{k,n} \frac{\delta_n^2}{2} + \sum_n \tau_{l,n} \frac{(\delta'_n)^2}{2} \right] .$$

Solving this constrained optimization problem yields the following solution

$$\delta_n = \frac{(\mu'_{l,m} - \mu'_{k,m})w_{n,m}/\tau_{k,n}}{\sum_{n' \in [M]} w_{n',m}^2 (1/\tau_{k,n'} + 1/\tau_{l,n'})} \text{ and } \delta'_n = \frac{(\mu'_{l,m} - \mu'_{k,m})w_{n,m}/\tau_{l,n}}{\sum_{n' \in [M]} w_{n',m}^2 (1/\tau_{k,n'} + 1/\tau_{l,n'})}.$$

We conclude similarly to the best arm identification case. \square

Note that we retrieve the same bound as for the case $N = 1$ (i.e., best arm identification).

Relaxed Lower Bound Problem for Top- N Identification For Top- N identification, let us define

$$\forall k, m, \Delta'_{k,m} := \begin{cases} \max_{k' \in [K]}^N \mu'_{k',m} - \mu'_{k,m} & \text{if } k \notin S_m^* \\ \mu'_{k,m} - \max_{k' \in [K]}^{N+1} \mu'_{k',m} & \text{otherwise,} \end{cases}$$

and $\tilde{N}_W^*(\mu)$ the value of problem $\tilde{\mathcal{P}}^* \left(\left((\Delta'_{k,m})^2 / 2 \right)_{k,m} \right)$. The set of constraints $\mathcal{N} := \left\{ t \mid \forall m \in [M], \forall k \notin S_m^*, \forall l \in S_m^*, \sum_n w_{n,m}^2 \left(\frac{1}{t_{k,n}} + \frac{1}{t_{l,n}} \right) \leq \frac{(\mu'_{k,m} - \mu'_{l,m})^2}{2} \right\}$ is included in the set of constraints $\tilde{\mathcal{N}} := \left\{ t \mid \forall m \in [M], \forall k \in [K], \sum_n \frac{w_{n,m}^2}{t_{k,n}} \leq \frac{(\Delta'_{k,m})^2}{2} \right\}$: indeed, if $t \in \mathcal{N}$, then for any $m \in [M]$, and any $k \notin S_m^*$,

$$\begin{aligned} \forall l \in S_m^*, \sum_n \frac{w_{n,m}^2}{t_{k,n}} &\leq \sum_n w_{n,m}^2 \left(\frac{1}{t_{k,n}} + \frac{1}{t_{l,n}} \right) \leq \frac{(\mu'_{k,m} - \mu'_{l,m})^2}{2} \\ \implies \sum_n \frac{w_{n,m}^2}{t_{k,n}} &\leq \min_{l \in S_m^*} \frac{(\mu'_{k,m} - \mu'_{l,m})^2}{2} = \frac{(\mu'_{k,m} - \max_{k' \in [K]}^N \mu'_{k',m})^2}{2} = \frac{(\Delta'_{k,m})^2}{2}. \end{aligned}$$

Similarly, for any agent m and $l \in S_m^*(\mu)$, one can check that $\sum_n \frac{w_{n,m}^2}{t_{l,n}} \leq \frac{(\Delta'_{l,m})^2}{2}$, hence $t \in \tilde{\mathcal{N}}$.

Then $N_W^*(\mu) \geq \tilde{N}_W^*(\mu)$.

Algorithm for Top- N identification Algorithm 1 can then easily be adapted to Top- N identification, with the following changes (the full algorithm is described in Algorithm 2)

1. Replace the stopping criterion (and the condition for the update of proxy gaps $(\tilde{\Delta}_{k,m}(r))_{k,m}$ at round r) with

$$\forall m \in [M], |B_m(r)| \leq N,$$

2. Replace the elimination criterion with

$$B_m(r+1) \leftarrow \left\{ k \in B_m(r) \mid \mu'_{k,m} + \Omega_{k,m}(r) \geq \max_{i \in B_m(r)}^N (\mu'_{i,m} - \Omega_{i,m}(r)) \right\}.$$

Remark 2. Note that the proxy gap $\tilde{\Delta}_{k,m}(r)$ no longer tracks the value of gap $\Delta'_{k,m}$ for $m \in [M], k \in [K]$, but $\Delta'_{k,m}$, and that on $N = 1$, this algorithm exactly coincides with Algorithm 1.

Analysis of Algorithm 2. First, such an algorithm is indeed δ -correct on event \mathcal{E} (the same as defined for Algorithm 1). Otherwise, for some agent m , there would be an arm $l \in S_m^*$ which is eliminated at round r from $B_m(r+1)$. But, on event \mathcal{E} , Lemma 3 implies that, for any $r \geq 0$, $m \in [M]$, and $(i, j) \in [K]^2$,

$$\hat{\mu}'_{i,m}(r) - \hat{\mu}'_{j,m}(r) + \Omega_{i,m}(r) + \Omega_{j,m}(r) \geq \mu'_{i,m} - \mu'_{j,m} \geq \hat{\mu}'_{i,m}(r) - \hat{\mu}'_{j,m}(r) - \Omega_{i,m}(r) - \Omega_{j,m}(r).$$

Algorithm 2 Weighted Collaborative Phased Elimination for Top- N identification (W-CPE-Top N)

Input: $\delta \in (0, 1)$, M agents, K arms, matrix W , $N \in [K]$

Initialize $r \leftarrow 0$, $\forall k, m, \tilde{\Delta}_{k,m}(0) \leftarrow 1, n_{k,m}(0) \leftarrow 1, \forall m, B_m(0) \leftarrow [K]$

Draw each arm k by each agent m once

repeat

Central server

$B(r) \leftarrow \cup_{m \in [M]} B_m(r)$

Compute $t(r) \leftarrow \tilde{\mathcal{P}}^* \left(\left(\sqrt{2} \tilde{\Delta}_{k,m}(r) \right)_{k,m} \right)$

For all $k \in [K]$, compute

$$(d_{k,m}(r))_{m \in [M]} \leftarrow \arg \min_{d \in \mathbb{N}^M} \sum_m d_m \text{ s.t. } \forall m \in [M], \frac{n_{k,m}(r-1) + d_m}{\beta_\delta(n_{k,\cdot}(r-1) + d)} \geq t_{k,m}(r)$$

Send to each agent m $(d_{k,m}(r))_{k,m}$ and $d_{\max} := \max_{n \in [M]} \sum_{k \in [K]} d_{k,n}(r)$

Agent m

Sample arm $k \in B(r)$ $d_{k,m}(r)$ times, so that $n_{k,m}(r) = n_{k,m}(r-1) + d_{k,m}(r)$

Remain idle for $d_{\max} - \sum_{k \in [K]} d_{k,m}(r)$ rounds

Send to the server empirical mean $\hat{\mu}_{k,m}(r) := \sum_{s \leq n_{k,m}(r)} X_{k,m}(s) / n_{k,m}(r)$ for any $k \in [K]$

Central server

Compute the empirical mixed means $(\hat{\mu}'_{k,m}(r))_{k,m}$ based on $(\hat{\mu}_{k,m}(r))_{k,m}$ and W

// Update set of candidate best arms for each user

for $m = 1$ **to** M **do**

$$B_m(r+1) \leftarrow \left\{ k \in B_m(r) \mid \hat{\mu}'_{k,m}(r) + \Omega_{k,m}(r) \geq \max_{j \in B_m(r)} \left(\hat{\mu}'_{j,m}(r) - \Omega_{j,m}(r) \right) \right\}$$

end for

// Update the gap estimates

For all $k, m, \tilde{\Delta}_{k,m}(r+1) \leftarrow \tilde{\Delta}_{k,m}(r) \times (1/2)^{\mathbb{1}(k \in B_m(r+1) \wedge |B_m(r+1)| > N)}$

$r \leftarrow r + 1$

until $\forall m \in [M], |B_m(r)| \leq N$

Output: $\{k \in B_m(r) : m \in [M]\}$

Then, combining the right-hand inequality for $j = l$ with the elimination criterion

$$\begin{aligned} \max_{i \in [K]} \mu'_{i,m} - \mu'_{l,m} &\geq \max_{i \in [K]} (\hat{\mu}'_{i,m}(r) - \hat{\mu}'_{l,m}(r) - \Omega_{i,m}(r) - \Omega_{l,m}(r)) \\ &\geq \max_{i \in B_m(r) \subseteq [K]} (\hat{\mu}'_{i,m}(r) - \hat{\mu}'_{l,m}(r) - \Omega_{i,m}(r) - \Omega_{l,m}(r)) > 0, \end{aligned}$$

which is absurd because $l \in S_m^*$. Then, let us consider the following notation, for any $m \in [M]$, $k \notin S_m^*$,

$$R_{k,m} := \sup\{r \geq 0 : k \in B_m(r)\} \text{ and } r_{k,m}^N := \min\{r \geq 0 : 4\tilde{\Delta}_{k,m}(r) < \Delta'_{k,m}{}^N\}.$$

Note that the random number of rounds used by Algorithm 2 is then $R^N = \max_{m \in [M]} \max_{k \notin S_m^*}^N R_{k,m}$. It is easy to prove that Lemma 4 and Lemma 10 still hold in Algorithm 2. An equivalent result to Lemma 11 can be shown

Lemma 17. *On event \mathcal{E} , for any $m \in [M]$, $k \notin S_m^*$, $R_{k,m} \leq r_{k,m}^N$.*

Proof. For any $m \in [M]$, $k \notin S_m^*$, $r = r_{k,m}$, if $k \notin B_m(r)$, then the claim is true. Otherwise, if $k \in B_m(r)$, then

$$\begin{aligned}
\hat{\mu}'_{k,m}(r) + \Omega_{k,m}(r) &\stackrel{(1)}{\leq} \mu'_{k,m} + 2\Omega_{k,m}(r) \\
&\stackrel{(2)}{\leq} \mu'_{k,m} + 4\tilde{\Delta}_{k,m}(r) - 2\tilde{\Delta}_{k,m}(r) \\
&\stackrel{(3)}{<} \max_{i \in [K]}^N \mu'_{i,m} - 2\tilde{\Delta}_{k,m}(r) \stackrel{(4)}{=} \max_{i \in B_m(r)}^N \mu'_{i,m} - 2\tilde{\Delta}_{k,m}(r) \\
&\stackrel{(1)}{\leq} \max_{i \in B_m(r)}^N (\hat{\mu}'_{i,m}(r) - \Omega_{i,m}(r) + 2\Omega_{i,m}(r)) - 2\tilde{\Delta}_{k,m}(r)
\end{aligned}$$

Then

$$\begin{aligned}
\hat{\mu}'_{k,m}(r) + \Omega_{k,m}(r) &\stackrel{(2)}{\leq} \max_{i \in B_m(r)}^N (\hat{\mu}'_{i,m}(r) - \Omega_{i,m}(r) + 2\tilde{\Delta}_{i,m}(r)) - 2\tilde{\Delta}_{k,m}(r) \\
&\leq \max_{i \in B_m(r)}^N (\hat{\mu}'_{i,m}(r) - \Omega_{i,m}(r)) + 2 \max_{i \in B_m(r)}^N \tilde{\Delta}_{i,m}(r) - 2\tilde{\Delta}_{k,m}(r) \\
&\stackrel{(5)}{=} \max_{i \in B_m(r)}^N (\hat{\mu}'_{i,m}(r) - \Omega_{i,m}(r)) + 2 \cdot 2^{-r} - 2 \cdot 2^{-r}, \\
\implies \hat{\mu}'_{k,m}(r) + \Omega_{k,m}(r) &< \max_{i \in B_m(r)}^N (\hat{\mu}'_{i,m}(r) - \Omega_{i,m}(r)).
\end{aligned}$$

where (1) is using by using event \mathcal{E} ; (2) is using Lemma 4; (3) uses $r = r_{k,m}^N$ and $k \notin S_m^*$; (4) is using event \mathcal{E} , and, for all $l \in S_m^*$, $l \in B_m(r)$; (5) holds because Lemma 10 is still valid and then, for all $j \in B_m(r)$, $\tilde{\Delta}_{j,m}(r) = 2^{-r}$. Then k is eliminated from $B_m(r)$ at round at most $r = r_{k,m}^N$, hence $R_{k,m} \leq r_{k,m}^N$. \square

Using this result, the sample complexity analysis is the same as for best arm identification, which yields

Theorem 5. *With probability $1 - \delta$, Algorithm 2 outputs the Top- N arms for each agent using a total number of samples no greater than*

$$\sup\{n \in \mathbb{N}^* : n \leq 32N_W^*(\mu) \log_2(8/\Delta'_{\min})\beta^*(n) + KM\},$$

where $\beta^*(n) := \beta_\delta(n\mathbb{1}_{[M]})$.

and we can use the same tools as in best arm identification to get an explicit upper bound depending on $N_W^*(\mu)$.

E Experimental study

The general weighted collaboration bandit framework has not been studied prior to this work. We investigate its performance in the special case of federated learning with personalization [31], which corresponds to choosing weight matrices of the form $w_{n,m} = \alpha\mathbb{1}(n=m) + \frac{1-\alpha}{M}$ for any pair of agents (n,m) . In this special case, we propose a baseline for weighted collaborative best arm identification which is a natural counterpart of the regret algorithm proposed by [31], and compare it to our W-CPE-BAI algorithm.

E.1 A Simple BAI Algorithm Inspired by PF-UCB

We state below as Algorithm 3 a straightforward adaptation of the PF-UCB algorithm in [30] to *personalized* federated best arm identification (BAI); meaning that only weight matrices of the form $w_{n,m} = \alpha\mathbb{1}(n=m) + \frac{1-\alpha}{M}$ for any pair of agents (n,m) are considered. The original regret algorithm uses phased eliminations designed for each agent to identify their best arm together with *exploitation*: when all best arms have been found, or when some agent is waiting for others to finish their own exploration rounds, agents keep playing their empirical best arm. To turn this into a δ -correct BAI algorithm, we remove the exploration rounds; keep the same sampling rule within each

phase (in which the number of samples from each arm is proportional to some rate function $f(r)$); and calibrate the size of the confidence intervals used to perform eliminations slightly differently, introducing for any $\delta \in (0, 1)$ function

$$\forall r \geq 0, B_r(\delta) := \sqrt{\frac{2 \log(KM\zeta(\beta)r^\beta/\delta)}{MF(r)}}.$$

where $F(r) = \sum_{p=1}^r f(p)$, for some $\beta > 1$. In practice, we use $\beta = 2$.

Algorithm 3, that we refer to as PF-UCB-BAI, follows the same general structure as our algorithm, with the notable difference that the number of samples of an arm $k \in B_m(r)$ in phase r is fixed in advance. Under PF-UCB-BAI, when arm k is still in the active set $B(r)$, agent m

- performs global exploration to sample it $d_{k,m}^g(r) := \lceil (1 - \alpha)f(r) \rceil$ times
- and additionally performs local exploration to sample it $d_{k,m}^\ell(r) := \lceil \alpha M f(r) \rceil$ extra times if furthermore $k \in B_m(r)$.

Overall, $d_{k,m}(r) := d_{k,m}^g(r) + d_{k,m}^\ell(r) = \lceil (1 - \alpha)f(r) \rceil \mathbb{1}(k \in B(r)) + \lceil \alpha M f(r) \rceil \mathbb{1}(k \in B_m(r))$ new samples from arm k are collected by agent m during phase r in order to update its estimate $\hat{\mu}_{k,m}(r)$ —the average of all available $n_{k,m}(r)$ samples for arm k obtained by agent m —which is sent to the central server.

The mixed mean of each arm (k, m) can then be computed by the server as

$$\hat{\mu}'_{k,m}(r) := \left(\alpha + \frac{1 - \alpha}{M} \right) \hat{\mu}_{k,m}(r) + \frac{1 - \alpha}{M} \sum_{m \neq n} \hat{\mu}_{k,n}(r),$$

and sent back to each agent. We note that, in [30], they propose that the server computes the average $\hat{\mu}_k(r) := \frac{1}{M} \sum_{m=1}^M \hat{\mu}_{k,m}(r)$ across agents, and sends this value to each agent, who can then obtain $\hat{\mu}'_{k,m}(r) := \alpha \hat{\mu}_{k,m}(r) + (1 - \alpha) \hat{\mu}_k(r)$.

Arm k is eliminated from the active set $B_m(r)$ of agent m if

$$\hat{\mu}'_{k,m}(r) + B_r(\delta) < \max_{j \in B_m(r)} (\hat{\mu}'_{j,m}(r) - B_p(\delta))$$

for the confidence parameter $B_r(\delta)$. In the original algorithm, $B_r(\delta)$ is replaced by some function of r and T , however a simple adaptation of Lemma 1 in [30] (adding a union bound on $r \in \mathbb{N}$) yields the following result. Indeed, the original result crucially exploits the sampling rule, which we did not change.

Lemma 18. *Event*

$$\mathcal{G} := \left\{ \forall r \in \mathbb{N}^*, \forall m \in [M], \forall k \in B_m(r), |\hat{\mu}'_{k,m}(r) - \mu'_{k,m}| \leq B_r(\delta) \right\}$$

holds with probability $1 - \delta$.

On the good event \mathcal{G} introduced in Lemma 18, observe that arm k_m^* can never be eliminated from the set $B_m(r)$, therefore it has to be the guess \hat{k}_m that agent m outputs. This proves that Algorithm 3 is δ -correct for pure exploration for the special case of federated bandit with personalization. This algorithm can therefore serve as a baseline to be compared to our proposal in this particular case.

We can also upper bound the sample complexity of this algorithm. Indeed, on event \mathcal{G} , like in the analysis of PF-UCB in [30], we can upper bound the number of rounds where arm k is sampled by agent m by $p_{k,m} := \inf\{r : B_r(\delta) \leq \Delta'_{k,m}/4\}$. When $f(p) = 2^p$, one can prove that

$$\sum_{p=1}^{p_{k,m}} f(p) = \mathcal{O} \left(\frac{\log(1/\delta)}{M(\Delta'_{k,m})^2} \right).$$

Summing the (deterministic) global and local exploration cost over rounds, arms and agents, yields an exploration cost of order

$$\mathcal{O} \left(\sum_{k \in [K]} \left[\left(\frac{1 - \alpha}{\min_{n \in [M]} (\Delta'_{k,m})^2} \right) + \left(\sum_{m \in [M]} \frac{\alpha}{(\Delta'_{k,m})^2} \right) \right] \log \left(\frac{1}{\delta} \right) \right).$$

Algorithm 3 PF-UCB-BAI

Input: $\delta \in (0, 1)$, M agents, K arms, matrix W .
 $f(r)$: sampling effort in phase r , $B_r(\delta)$: size of the confidence intervals in phase r .
Initialize $r \leftarrow 0$, $\forall k, m, n_{k,m}(0) \leftarrow 0$, $\forall m, B_m(0) \leftarrow [K]$, $\hat{k}_m \leftarrow 0$.

repeat

- # Central server
- if** $|B_m(r)| = 1$ **then**
 - $\hat{k}_m \leftarrow$ the unique arm in $B_m(r)$
 - $B_m(r) = \emptyset$
- end if**
- $B(r) \leftarrow \bigcup_{m \in [M]} B_m(r)$
- for** $k \in B(r), m \in [M]$ **do**
 - $d_{k,m}(r) = \lceil (1 - \alpha)f(r) \rceil + \lceil \alpha M f(r) \rceil \mathbb{1}(k \in B_m(r))$
- end for**
- Send to each agent m $(d_{k,m}(r))_{k,m}$ and $d_{\max} := \max_{n \in [M]} \sum_{k \in [K]} d_{k,n}(r)$

- # Agent m
- Sample arm $k \in B(r)$ $d_{k,m}(r)$ times, so that $n_{k,m}(r) = n_{k,m}(r-1) + d_{k,m}(r)$
- Remain idle for $d_{\max} - \sum_{k \in [K]} d_{k,m}(r)$ rounds
- Send to the server empirical mean $\hat{\mu}_{k,m}(r) := \sum_{s \leq n_{k,m}(r)} X_{k,m}(s) / n_{k,m}(r)$ for any $k \in [K]$

- # Central server
- Compute the empirical mixed means $(\hat{\mu}'_{k,m}(r))_{k,m}$ based on $(\hat{\mu}_{k,m}(r))_{k,m}$ and W
- // Update set of candidate best arms for each user
- for** $m = 1$ **to** M **do**
 - $B_m(r+1) \leftarrow \left\{ k \in B_m(r) \mid \hat{\mu}'_{k,m}(r) + B_r(\delta) \geq \max_{j \in B_m(r)} (\hat{\mu}'_{j,m}(r) - B_r(\delta)) \right\}$
- end for**
- $r \leftarrow r + 1$
- until** $|B(r)| = \emptyset$
- Output:** $\{\hat{k}_m : m \in [M]\}$

E.2 Numerical experiments

As we did throughout the paper, we consider Gaussian bandits with fixed variance $\sigma^2 = 1$. We denote \hat{r} the average number of communication rounds across the R iterations of an experiment ; \hat{c} the average exploration cost of the considered algorithm across the R iterations ; $\hat{\delta}$ the empirical error frequency across the R iterations. \hat{r} and \hat{c} are reported \pm their standard deviation rounded up to the closest integer, except for $\hat{\delta}$, which is rounded up to the 2nd decimal place.

We consider a synthetic instance with $K = 6$ arms, $M = 3$ agents, for $R = 100$ iterations. In order to generate randomly this instance, we sampled at random $K \times M$ values $x_{k,m}$ from the distribution $\mathcal{N}(0, 1)$, and set $\mu_{k,m} = x_{k,m} / \|x\|_{\mathcal{F}}$,⁹ and tested if the associated Δ'_{\min} satisfied $\Delta'_{\min} \geq 0.05$. We repeated this sampling until this condition was fulfilled.

Comparison to PF-UCB-BAI Our first experiment is to compare our Algorithm 1 with the PF-UCB-BAI baseline described above. For PF-UCB-BAI we use a phase length $f(p) := 2^p \log(1/\delta)$ for $p \geq 0$. For both algorithms, we set $\delta = 0.1$ and experiment with $\alpha \in \{0.4, 0.5, 0.6, 0.7\}$. Results are reported on the left-most table in Table 1. In terms of communication cost, W-CPE-BAI (Algorithm 1) improves considerably over the baseline. Depending on the value $\alpha \in [0, 1]$ (the closer it is to 1, the less agents have to communicate in order to get good estimates of their mixed expected reward) W-CPE-BAI improves or has an exploration cost which is similar to the baseline, up to a constant lower than 2.

⁹ $\|\cdot\|_{\mathcal{F}}$ is the Frobenius norm for matrices.

Table 1: Personalized collaborative BAI with varying $\alpha \in \{0.4, 0.5, 0.6, 0.7\}$; the most efficient algorithm in terms of exploration cost is in bold type , all algorithms yield $\hat{\delta} = 0$ up to the 5th decimal place (*top table*). Personalized collaborative BAI with varying $\delta \in \{0.00001, 0.0001, 0.001, 0.01, 0.05, 0.1\}$; ratios are rounded up to the 1st decimal place.

α	ALGORITHM	\hat{r}	\hat{c}	$\hat{\delta}$	δ	\hat{c}	c^*	\hat{c}/c^*
0.4	W-CPE-BAI	5± 0	87,918± 8,634	0.00	0.1 0.05 0.01 0.001 0.0001 0.00001	75,094± 7,596 75,992± 8,468 80,457± 8,838 110,888± 31,613 91,692± 9,121 96,942± 9,885	1,104 1,639 2,875 4,643 6,411 8,180	68.0 46.4 28.0 23.9 14.3 11.9
	PF-UCB-BAI	12± 0	109,822± 30,335	0.00				
0.5	W-CPE-BAI	5± 0	75,094± 7,596	0.00				
	PF-UCB-BAI	11± 0	73,561± 17,239	0.00				
0.6	W-CPE-BAI	5± 0	78,334± 7,983	0.00				
	PF-UCB-BAI	11± 0	55,812± 14,793	0.00				
0.7	W-CPE-BAI	5± 0	76,817± 10,953	0.00				
	PF-UCB-BAI	10± 0	45,765± 9,591	0.00				

Comparison to an oracle algorithm Our second experiment is to assess the asymptotic optimality (up to some logarithmic factors) of our algorithm. In order to estimate the scaling in $T^*(\cdot)$, we have implemented an oracle algorithm which has access to the true gaps $\left((\Delta'_{k,m})_{k,m} \right)$ and can compute the associated $T_W^*(\mu)$. Then we compare the average exploration cost \hat{c} with $c^* := \left\lceil T_W^*(\mu) \log \left(\frac{1}{2.4\delta} \right) \right\rceil$, for $\delta \in \{0.00001, 0.0001, 0.001, 0.01, 0.05, 0.1\}$, and $\alpha = 0.5$. We reported the associated results in the right-most table in Table 1. We can notice that, as δ decreases, the ratio \hat{c}/c^* also decreases. As predicted by our upper bound in Theorem 2, W-CPE-BAI does not attain asymptotic optimality even for small values of δ , but has a scaling to $T_W^*(\mu) \log \left(\frac{1}{2.4\delta} \right)$ which decreases as δ goes to 0.

Numerical considerations These experiments were run on a personal computer (configuration: processor Intel Core i7-8750H, 12 cores @2.20GHz, RAM 16GB). In order to solve the optimization problem defining the oracle $t(r)$, we used CVXPY [9, 1], with the commercial solver MOSEK [27] tuned to default parameters. To compute the number of samples $(d_{k,m}(r))_{k,m,r}$, we used the *optimize* module from SciPy [33]. In our experiments with this implementation of W-CPE-BAI, we found that when the instance is too hard –meaning that the associated Δ'_{\min} is small– the optimization part is subject to numerical approximation errors, which prevents the computation of the oracle allocation. This might however be mitigated by online optimization approaches, such as in [8].