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## Occupational exposure to nickel and hexavalent chromium and the risk of lung cancer in a pooled analysis of case-control studies (SYNERGY)

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## List of abbreviations

$\mathrm{Cl} \quad$ Confidence interval

Cr (III) Trivalent chromium
$\mathrm{Cr}(\mathrm{VI}) \quad$ Hexavalent chromium
IARC International Agency for Research on Cancer
LOD Limit of detection
OR Odds ratio
RERI Relative excess risk due to interaction
SCOEL European Union Scientific Committee on Occupational Exposure Limits

## Novelty and Impact

Assessment of lung-cancer risk after occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel has been primarily investigated in industry-based studies so far.

SYNERGY utilized a population-based approach to study quantitative exposure-effect relationships using secondary measurement data from various regions and time periods across a wide range of jobs, while adjusting for smoking habits.

The observed increased lung-cancer risks at low cumulative exposure levels warrant continuing awareness to monitor the impact of occupational metal exposure on human cancer.


#### Abstract

There is limited evidence regarding the exposure-effect relationship between lung-cancer risk and hexavalent chromium $(\mathrm{Cr}(\mathrm{VI}))$ or nickel. We estimated lung-cancer risks in relation to quantitative indices of occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel and their interaction with smoking habits.

We pooled 14 case-control studies from Europe and Canada, including 16,901 lung-cancer cases and 20,965 control subjects. A measurement-based job-exposure-matrix estimated job-yearregion specific exposure levels to $\mathrm{Cr}(\mathrm{VI})$ and nickel, which were linked to the subjects' occupational histories. Odds ratios (OR) and associated 95\% confidence intervals (CI) were calculated by unconditional logistic regression, adjusting for study, age group, smoking habits, and exposure to other occupational lung carcinogens. Due to their high correlation, we refrained from mutually adjusting for $\mathrm{Cr}(\mathrm{VI})$ and nickel independently.

In men, ORs for the highest quartile of cumulative exposure to $\mathrm{CR}(\mathrm{VI})$ were $1.32(95 \% \mathrm{CI} 1.19$ $1.47)$ and 1.29 ( $95 \% \mathrm{Cl} 1.15-1.45$ ) in relation to nickel. Analogous results among women were: 1.04 ( $95 \% \mathrm{Cl} 0.48-2.24$ ) and 1.29 ( $95 \% \mathrm{Cl} 0.60-2.86$ ), respectively. In men, excess lung-cancer risks due to occupational $\mathrm{Cr}(\mathrm{VI})$ and nickel exposure were also observed in each stratum of never, former and current smokers. Joint effects of $\mathrm{Cr}(\mathrm{VI})$ and nickel with smoking were in general greater than additive, but not differerent from multiplicative.

In summary, relatively low cumulative levels of occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel were associated with increased ORs for lung cancer, particularly in men. However, we cannot rule out a combined classical measurement and Berkson-type of error structure, which may cause differential bias of risk estimates.


## INTRODUCTION

The hexavalent form of chromium $(\mathrm{Cr}(\mathrm{VI}))$ has been long recognized as human carcinogen. ${ }^{1}$ Exposure mainly arises in hot metal processes, during the processing of stainless steel, during surface treatment by polishing, sanding, and grinding, and, historically, during the manufacture of chromium pigment. ${ }^{2,3}$ In previous analyses of the German MEGA measurement database, we observed the highest $\mathrm{Cr}(\mathrm{VI})$ concentrations in spray painting and hard-chromium plating, and also in welding fumes from shielded metal arc welding and flux-cored arc welding. ${ }^{4}$ Determination of $\mathrm{Cr}(\mathrm{VI})$ is difficult as it is frequently deoxidized to the more stable trivalent chromium $(\mathrm{Cr}(\mathrm{III})) .{ }^{5}$ In contrast to $\mathrm{Cr}(\mathrm{III}), \mathrm{Cr}(\mathrm{VI})$ readily passes cell membranes. Intracellular reduction to $\mathrm{Cr}(\mathrm{III})$ may lead to oxidative stress, resulting in protein and DNA damage, genomic instability, cytotoxicity, tissue damage, chronic inflammation, and epigenetic changes such as microRNA, histone modification, and DNA methylation which all may trigger carcinogenesis. ${ }^{6}$ The European Union Scientific Committee on Occupational Exposure Limits (SCOEL) estimated an "acceptable" lifetime excess risk of four additional lung-cancer cases per 1,000 after a 40-year occupational exposure to 1 $\mu \mathrm{g} / \mathrm{m}^{3}$ of $\mathrm{Cr}(\mathrm{VI})\left(40 \mu \mathrm{~g} / \mathrm{m}^{3}\right.$-years $) .{ }^{7}$

Nickel is a widespread occupational exposure in various jobs and industries, frequently with coexposure to chromium. ${ }^{1}$ Exposure frequently occurs in nickel alloy and battery production. ${ }^{8}$ It has been demonstrated that workers in several industrial processes (e.g. metal-cutting and metalforming activities, metal spraying, sintering, chemical production, manufacturing of glass, batteries and accumulators, as well as certain welding processes) have experienced exposures at median nickel concentrations above $10 \mu \mathrm{~g} / \mathrm{m}^{3}$, which is the recommended SCOEL threshold limit value to protect workers from carcinogenicity. ${ }^{9,10}$ As early as 1979, working in nickel refineries was classified as Group 1 carcinogen by the International Agency for Research on Cancer (IARC), ${ }^{11}$ and the same classification was later also assigned to various nickel compounds. ${ }^{1}$

So far, epidemiological evidence on the exposure-effect relationship between occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ or nickel with lung-cancer risk primarily has been obtained from studies among workers in chromate production and in nickel refining. ${ }^{12-14}$ Increased lung-cancer risks were also described in chromate pigment production and among chrome plating workers. ${ }^{2}$

At-risk occupations with exposures to $\mathrm{Cr}(\mathrm{VI})$ and nickel comprise welders as the largest workforce. Welding fumes have been classified as a Group 1 carcinogen, ${ }^{15}$ and several job titlebased analyses have demonstrated increased risks for lung cancer among professional, ${ }^{16-18}$ but also occasional welders. ${ }^{18}$ Due to the complex composition of welding fumes, it is challenging to attribute lung-cancer risk to one of its major components, which may be illustrated by the inability of many studies to demonstrate consistently elevated lung-cancer risks to $\operatorname{Cr}(\mathrm{VI})$ or nickel exposure in association with welding activities. ${ }^{16,17}$

There is little evidence showing quantitative, measurement-based exposure-effect relationships between $\mathrm{Cr}(\mathrm{VI})$ or nickel and lung cancer across a wide array of job activities, while adjusting for smoking habits. We therefore took advantage of data from the pooled SYNERGY case-control study of occupational lung cancer to estimate relative risks related to occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ or nickel. The objectives of this paper were: 1 ) to estimate lung-cancer risk associated with quantitative indices of occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and to nickel; 2) to assess the shape of the exposure-effect relationship between each metal and lung cancer separately; and 3) to assess their joint effects with smoking habits.

## METHODS

## SYNERGY Project

The detailed objectives, methods, and aims of SYNERGY are described elsewhere. ${ }^{19,20}$ Briefly, SYNERGY was established as an international pooled case-control study to investigate joint
effects of occupational carcinogens (asbestos, ${ }^{19}$ respirable crystalline silica, ${ }^{20}$ polycyclic aromatic hydrocarbons, ${ }^{21}$ chromium, nickel) and smoking ${ }^{22}$ in the development of lung cancer. Over the years, this study has developed into an international platform for research on occupational lung cancer with 16 case-control studies from 22 study centers. For this analysis, we used data from the 14 original SYNERGY studies from Europe and Canada (Table S1), including 16,901 lungcancer cases and 20,965 control subjects. More information about SYNERGY is available at http://synergy.iarc.fr.

## Assessment of occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel

The development of the quantitative job-exposure matrix SYN-JEM to assess occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel followed a protocol which has been described in detail elsewhere. ${ }^{23}$ Briefly, personal measurements of chromium ( $n=24,150$ ) and nickel ( $n=22,081$ ), covering a period from the 1970s to 2009, were collected in the participating countries, compiled in the ExpoSYN database, and tagged with an ISCO-68 job title. Overall, $35 \%$ of the chromium and $28 \%$ of the nickel measurements were below the limit of detection (LOD). ${ }^{23}$ We substituted these measurements with a random figure between 0 and the LOD, assuming that they followed the same log-normal probability distribution as the measurements above LOD. ${ }^{24}$

A standard linear mixed-effects model was developed to assign region- and time-specific exposure levels for each ISCO-68-based job title that was solicited from the subjects' selfreported job histories. Region/country and job title were used as random effects, whereas year of measurement, sampling duration, and a prior exposure rating from a semi-quantitative expertbased job-exposure matrix (DOMJEM) assigning no, low, or high exposure levels ${ }^{25}$ were included as fixed effects. The DOMJEM rating was used as an override for non-routine measurements to set jobs considered to be non-exposed to $0 \mu \mathrm{~g} / \mathrm{m}^{3}$. In addition, models for $\operatorname{Cr}(\mathrm{VI})$ and nickel included the analytical method as fixed effect. Measurements conducted for jobs that were
assumed by SYN-JEM to be non-exposed were retained in the model for the overall assessment of time trends and regional differences in $\mathrm{Cr}(\mathrm{VI})$ and nickel levels. Model-based estimates were used to calculate the amount of $\operatorname{Cr}(\mathrm{VI})$ based on specific $\operatorname{Cr}(\mathrm{VI})(n=8,363)$ and total chromium measurements ( $n=15,787$ ). For total chromium values a conversion factor set at a total chromium: $\mathrm{Cr}(\mathrm{VI})$ ratio of $3: 1$ was applied. ${ }^{23}$

The model yielded a linear temporal trend with an annual decrease of $\mathrm{Cr}(\mathrm{VI})$ concentrations of $-2.7 \%$ and $-1.2 \%$ per year for nickel. When there were $<5$ measurements for a specific job, the mean estimate of all jobs within the same unit or major group was used to base a job-specific exposure estimate on information from similar jobs. ${ }^{23}$

It should be noted that assigning quantitative exposure data as part of a job-exposure matrix may lead to a combined classical measurement and Berkson-type of error structure, which may cause over- or underestimation of coefficients in logistic-regression analysis. ${ }^{26}$

Lifetime cumulative exposure to $\mathrm{Cr}(\mathrm{VI})$ or nickel was calculated as the sum of the country/regionspecific SYN-JEM estimates for each job and year. Cumulative indices were categorized according to quartiles based on the distribution among all (both sexes combined) control subjects. For interaction analyses a cutoff at the median was applied to define low and high exposure categories.

## Statistical analysis

We calculated odds ratios (OR) with $95 \%$ confidence intervals (CI) by unconditional logistic regression analysis. The main models included either $\mathrm{Cr}(\mathrm{VI})$ or nickel as the exposure variable, in addition to a number of covariates. Mutual adjustment was not performed in the main models, because a strong correlation between cumulative $\mathrm{Cr}(\mathrm{VI})$ and nickel levels was observed in subjects with co-exposure to both metals (Pearson $\mathrm{r}=0.75$; 95\% CI 0.74-0.76). The reference
category therefore consisted of subjects who were not exposed occupationally to either $\operatorname{Cr}(\mathrm{VI})$ or to nickel.

To control for confounding, we employed two different models: OR1 was adjusted for study and age group (<45, 45-49, 50-54, 55-59, 60-64, 65-69, 70-74, 75+ years) and OR2 was additionally adjusted for cumulative cigarette consumption (log(cigarette pack-years+1), smoking status including time-since-quitting smoking cigarettes (current smokers, stopping smoking 2-7 years, 8-15 years, 16-25, 26+ years before interview/diagnosis, never smokers), and ever employment in a "list A" job (yes/no). List A includes occupations and industries with an established lung cancer risk. ${ }^{27,28}$ This approach is consistent with the other analyses in SYNERGY. ${ }^{19-21}$

Cigarette pack-years were calculated as smoking duration (years) x average cigarette smoking intensity per day/20. Current smokers included smokers who had stopped smoking within the last two years before the interview/diagnosis. Never smokers were defined as lifelong non-smokers and subjects with a smoking history of $<1$ pack-year.

To visualize the functional form of the adjusted exposure-effect relationship between each agent ( $\mathrm{Cr}(\mathrm{VI})$ or nickel) and lung-cancer risk for the fully adjusted model (OR2), we estimated restricted cubic spline functions and associated $95 \% \mathrm{Cl}$. The optimal smoothing parameter was selected based on generalized cross-validation and under the assumption that the total number of degrees of freedom required for a biologically plausible model would not exceed three. Restricted cubic spline analyses also included lagged analyses, neglecting exposures that occurred $5,10,15$, or 20 years before diagnosis (cases) or the interview (control subjects).

We assessed the additive interaction between smoking and $\mathrm{Cr}(\mathrm{VI})$ and nickel by estimating the 'relative excess risk due to interaction' (RERI). ${ }^{29}$ Possible departure from a multiplicative joint effect was assessed by testing a multiplicative interaction term in the statistical model.

We conducted several subgroup and sensitivity analyses to assess the robustness of our results:
a) We stratified analyses by hospital-based and population-based studies and study region (Northern Europe (Germany, Sweden, France, UK, The Netherlands); Southern Europe (Italy, Spain); East Europe (Czech Republic, Hungary, Poland, Romania, Russia, Slovakia); and Canada). b) We restricted the study base to blue-collar workers to rule out a general blue-collar worker effect (i.e., an increased risk associated with multiple hazardous exposures in blue-collar job activities). c) We excluded welders and d) we restricted analyses to workers who started working in 1960 as well as 1970 or later, because exposure data were scarce before the 1970s. e) Although the main analyses contained only one of the two exposure variables of interest, we conducted a set of sensitivity analyses that included both $\mathrm{Cr}(\mathrm{VI})$ and nickel.

Statistical analyses were carried out using R statistical software (version 3.6.1).

## RESULTS

Among men, lifetime prevalence of exposure to $\mathrm{Cr}(\mathrm{VI})$ was $30 \%$ among cases and $23 \%$ among controls. Exposure prevalence to nickel was $24 \%$ (cases) and $19 \%$ (controls), of whom $77.7 \%$ of cases and $83 \%$ of controls were also exposed to $\mathrm{Cr}(\mathrm{VI})$. As expected, exposure prevalence was much lower among women than men ( $5 \%$ to $\mathrm{Cr}(\mathrm{VI})$ in both cases and controls). The exposure prevalence to nickel in females was $3 \%$ among both cases and controls (Table 2).

Differences in median cumulative exposure levels to $\mathrm{Cr}(\mathrm{VI})$ and nickel were less pronounced. The median $\operatorname{Cr}(\mathrm{VI})$ exposure in men was: $42.8 \mu \mathrm{~g} / \mathrm{m}^{3}$-years (cases) and $40.8 \mu \mathrm{~g} / \mathrm{m}^{3}$-years (controls) and in women $26.2 \mu \mathrm{~g} / \mathrm{m}^{3}$-years (cases) and $26 \mu \mathrm{~g} / \mathrm{m}^{3}$-years (controls). Median nickel exposure among men was $22.7 \mu \mathrm{~g} / \mathrm{m}^{3}$-years among cases and $21.5 \mu \mathrm{~g} / \mathrm{m}^{3}$-years among controls. Women showed $16.7 \mu \mathrm{~g} / \mathrm{m}^{3}$-years (cases) and $14.2 \mu \mathrm{~g} / \mathrm{m}^{3}$-years (controls), respectively (Table 1).

We observed similarly increased lung-cancer ORs for ever exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel among both sexes. Assessment of cumulative exposure revealed a a close to monotonic exposure-effect trend among men in the fully adjusted model (ORs for the highest exposure category: $\mathrm{Cr}(\mathrm{VI})$ : $>99.5 \mu \mathrm{~g} / \mathrm{m}^{3}$-years; $\mathrm{OR}=1.32,95 \% \mathrm{Cl} 1.19-1.47$ and nickel: $>78.1 \mu \mathrm{~g} / \mathrm{m}^{3}$-years, $\mathrm{OR}=1.29 ; 95 \% \mathrm{Cl}$ 1.15-1.45) (Table 2). For women, the exposure-effect relationships were less consistent with $\mathrm{OR}=1.04 ; 95 \% \mathrm{Cl} 0.48-2.24$ in the highest $\mathrm{Cr}(\mathrm{VI})$ category and $\mathrm{OR}=1.29 ; 95 \% \mathrm{CI} 0.60-2.86$ in the highest nickel-exposure category (Table 2).

In men, we also observed a monotonic trend towards higher risk estimates with increasing duration of exposure to $\mathrm{Cr}(\mathrm{VI})$. Exposure for 30 years and more, compared to never exposed, showed increased ORs in the fully adjusted model (OR=1.37; 95\% CI 1.23-1.51 for $\operatorname{Cr}(\mathrm{VI})$ and $\mathrm{OR}=1.23 ; 95 \% \mathrm{Cl}$ 1.09-1.38 for nickel). Risks peaked 10-19 years after cessation of exposure to $\mathrm{Cr}(\mathrm{VI})$ or nickel and then continuously declined towards baseline risk. The findings for women were less consistent (Table 2).

Subgroup analyses among males revealed more cases and slightly higher ORs in populationbased studies than hospital-based studies (Tables 3a and 3b). Compared with the full model, restricting the study base to male blue-collar workers and workers who started their job after 1960 showed a weaker exposure-effect relationship for $\operatorname{Cr}(\mathrm{VI})$ and nickel, although the highest exposure category still resulted in significantly increased ORs for lung cancer. Analyses restricted to workers starting after 1970 showed similar risk patterns, albeit less strong. Subgroup analyses among female subjects were based on few cases only, and the results were quite imprecise (Tables S3a and S3b). Lagging exposure by 5, 10, 15, and 20 years generated similar results (Table S4) compared to the unlagged risk estimates.

Center-specific results revealed some heterogeneity between study regions where results from Southern Europe matched those from the North European region showing increased ORs,
whereas the picture was less homogeneous and the number of cases smaller in the other geographically similar study centers (see Tables S2a and S2b and Figures S1a and S1b).

Analyses using cubic splines showed a nearly linear exposure-effect relationship for nickel among males. The exposure-effect for $\mathrm{Cr}(\mathrm{VI})$ among males and among female subjects for both metals were linear (Figure 1).

When we stratified the analyses by smoking status, we observed similarly increased ORs for $\mathrm{Cr}(\mathrm{VI})$ exposure above the median ( $40.23 \mu \mathrm{~g} / \mathrm{m}^{3}$-years) among current smokers, former smokers, and never smokers (Tables S5-S7). In men, ORs were strongest for squamous-cell and smallcell lung cancer subtypes, whereas there was no consistent association between $\mathrm{Cr}(\mathrm{VI})$ exposure and adenocarcinoma. Risk estimates in never smokers appeared to be strongest for small-cell cancer of the lung (Table S5). Similar patterns were observed for occupational nickel exposure above the median of $30.75 \mu \mathrm{~g} / \mathrm{m}^{3}$-years (Table S6).

Among men, the joint effect of smoking and $\mathrm{Cr}(\mathrm{VI})$ on the risk of all lung-cancer subtypes was larger than additive (RERI=2.10; 95\% CI 1.41-2.79; Table 4). RERI was particularly high for squamous-cell and small-cell cancer of the lung (supplementary table S8). For women, these associations were similar, however estimates were less precise compared to men (Tables 4 and S8). When using a multiplicative model as framework, no statistical siginificance for the interaction term in the model was observed except for small-cell lung cancer in men, implying that there is no significant deviation of the joint effect between smoking and occupational exposure to $\operatorname{Cr}(\mathrm{VI})$ or nickel from multiplicativity (Tables 4 and S 8 ). Analysis of the interaction between $\mathrm{Cr}(\mathrm{VI})$ and nickel exposure was impaired by a high correlation between these agents, and, as stated above, we refrained from adjusting mutually for the other metal in these analyses.

Although the main analyses contained only one of the two exposure variables of interest, we conducted a set of sensitivity analyses that mutually adjusted for both, $\mathrm{Cr}(\mathrm{VI})$ and nickel. The OR
for $\mathrm{Cr}(\mathrm{VI})$ was similar in the two-variable model compared to the one-variable model. By contrast there was some difference between the two models for the OR estimate for nickel. In men, ORs for nickel were attenuated to $\mathrm{OR}=1.05 ; 95 \% \mathrm{CI} 0.97-1.14$ and ORs for women slightly increased (OR=1.39; 95\% Cl 0.99-1.94). Analysis of subjects solely exposed to $\mathrm{Cr}(\mathrm{VI})$, but not nickel yielded an OR of $1.40 ; 95 \% \mathrm{Cl} 1.25-1.57$ for men and $0.59 ; 95 \% \mathrm{CI} 0.24-1.42$ for women, but the latter analysis was based on only 10 exposed cases and 15 exposed controls. Subjects solely exposed to nickel were too few to conduct a sound sensitivity analysis (four male cases and three controls, but no exposed female case subject, all results not shown).

## DISCUSSION

We studied the associations between occupational $\mathrm{Cr}(\mathrm{VI})$ and nickel exposure with lung cancer in the pooled SYNERGY case-control study. Increasing duration and increasing cumulative exposure to $\mathrm{Cr}(\mathrm{VI})$ or nickel were associated with increasing ORs for lung cancer. As it can be expected from welding and various metalwork-related activities, ${ }^{1} \mathrm{Cr}(\mathrm{VI})$ and nickel exposures were highly correlated in our data so that we did not adjust mutually for both metals in our analyses. Increased risks for $\mathrm{Cr}(\mathrm{VI})$ and nickel were found in never smokers, former smokers, and current smokers. The joint effect of smoking and $\operatorname{Cr}(\mathrm{VI})$ or nickel exposure was generally more than additive, particularly for squamous-cell and small-cell cancer of the lung. All these effects were clearly seen in men with narrow confidence intervals. Women showed similar risks, but analyses were limited by smaller numbers of exposed subjects, and subsequently analyses yielded wider confidence intervals in ever-never comparisons and exposure-effect trends.

Hexavalent chromium and at least some forms of nickel compounds are established lung carcinogens which have been repeatedly evaluated by IARC. ., 2, 11, 30, 31 IARC's classification as

Group 1 carcinogens relied mainly on industrial cohort studies of chromium production and nickel refinery workers.

Two of the largest chromium cohorts from Baltimore, MD, and Painsville, OH , have been repeatedly updated and re-analyzed with respect to lung-cancer risk. These studies unanimously indicated some increase in lung-cancer risk with respect to occupational $\mathrm{Cr}(\mathrm{VI})$ exposure. ${ }^{12,32-35}$ We here add to the evidence by supporting these observations with analyses in a large pooled case-control study.

Although the exact nickel compounds responsible for an increased lung-cancer risk are unknown, results of studies among Norwegian refinery workers, suggest the strongest evidence for total nickel and water-soluble nickel compounds. ${ }^{2,14,36}$ Additional analyses of this Norwegian cohort revealed a clear dose-effect relationship with lung-cancer risk for water-soluble compounds, but little support for metallic, oxidic, or sulfidic forms of nickel as risk factors, when mutually adjusting for water-soluble nickel compounds. ${ }^{36}$ Although epidemiological studies are limited in disentangling, which form is associated with an increased lung-cancer risk due to exposure to multiple forms of nickel, our findings in SYNERGY compare well with findings from these cohorts (see OR1 in table 2). However, exposure levels were in general lower than in these industries.

More recently, a semi-quantitative approach was undertaken by a population-based Canadian case-control study, whose data partially also contributed to this analysis. The study also showed increased lung-cancer risks in relation to occupational $\mathrm{Cr}(\mathrm{VI})$ or nickel exposure, but only among non-smokers and former smokers quitting smoking over 20 years prior to inclusion into the study. ${ }^{37}$ This finding is probably due to the strong effects of smoking on lung-cancer risk, leading to relative risks for occupational exposures being superimposed by the higher risk for lung cancer from smoking.

The median cumulative exposure level of $40 \mu \mathrm{~g} / \mathrm{m}^{3}$-years for $\mathrm{Cr}(\mathrm{VI})$ observed in this study corresponds with the current SCOEL benchmark value of $1 \mu \mathrm{~g} / \mathrm{m}^{3}$ associated with $4 / 1,000$ excess lung-cancer cases during 40 years of working life, ${ }^{7}$ indicating that in the past a substantial part of the occupational workforce was exposed to $\mathrm{Cr}(\mathrm{VI})$ above these levels. The SCOEL benchmark value ${ }^{7}$ was based on the mean value from the individual slope estimates of $\beta=1.75,{ }^{38}$ as derived from the studies by Crump and co-workers $(\beta=0.68)^{12}$ and Park et al. ( $\beta=2.82$ ). ${ }^{33}$ Kauermann and others, ${ }^{39}$ using a variety of model specifications, derived a combined slope estimate of 0.63 based on a pooled analysis of aggregated data from these studies. Using these two reported slope estimates, we can calculate an expected relative risk of 1.19 and 1.07 for men at an exposure level of $0.1 \mathrm{mg} / \mathrm{m}^{3}$-years, respectively, which is in line with our finding of an OR of 1.24 at this exposure level. In contrast, median cumulative nickel concentrations in SYNERGY $\left(20 \mu \mathrm{~g} / \mathrm{m}^{3}\right.$ years) were much lower than the current SCOEL threshold limit value of $10 \mu \mathrm{~g} / \mathrm{m}^{3}$, if taking into account a 40-year occupational exposure period.

Strengths of our analysis include a large study population with sufficient power to detect potentially increased risks in subgroups such as women, non-smokers, and for histological lungcancer subtypes, while taking into account detailed information on smoking habits. The use of a database of measurements from different countries and industries and modelling of an exposure time trend enabled us to assess cumulative exposures over the entire job histories and across jobs and industries quantitatively. ${ }^{23}$

Although we included a high number of personal measurements to assess occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel, limitations related to exposure assessment are that the measurements for a particular job did not necessarily correspond with the jobs reported in the study subjects' occupational history. ${ }^{23}$ This will cause some degree of measurement error of the Berkson type, which will primarily weaken the precision of our estimates. Likely, the effect on point estimates will be rather small and lead to attenuation of ORs. ${ }^{40}$ However, we cannot rule out the possibility
of a combined error structure of classical measurement and Berkson-type of error, which may occur when estimating quantitative exposure-effect associations using random exposuregrouping methods. This is frequently the case in job-exposure matrices in which exposure levels are estimated for various occupational groups instead a fixed occupational setting. This situation may cause a non-differential measurement error turning into differential bias, thus leading to overor underestimation of risk estimates. As it has been shown, decreasing between-group variance usually leads to an increase in bias, which may also have affected our estimates that were situated in the low-exposure range. ${ }^{26}$

Exposure assessment in SYNERGY was performed to capture a wide array of exposed jobs, which may have resulted in assigning exposure levels to subjects who were only occasionally exposed to $\mathrm{Cr}(\mathrm{VI})$ or nickel or not at all. We therefore cannot rule out that our risk estimates, at least partially, entail a "blue-collar" effect associated with multiple exposures to several coccupational carcinogens. Restricting analyses to blue-collar workers, indeed revealed reduced ORs compared to the full sample. In addition, the relatively low response proportions in many of the population-based case-control studies may have resulted in a general under-representation of blue-collar workers in the control group, potentially inflating the observed associations when including white-collar workers. However, positive associations were seen for the highest exposure groups, and trends across exposure categories were consistent.

## CONCLUSIONS

Summarizing, we observed positive exposure-effect associations between lung-cancer risk and occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and to nickel in the large SYNERGY study in men. Women showed similar tendencies, albeit with less statistical precision due to the smaller numbers of exposed female subjects. We estimated exposure-risk relationships over a wide range of exposed
jobs, using a comprehensive measurement-based JEM. Among men, increased lung-cancer risks were associated with both longer exposure duration and higher cumulative exposure to $\mathrm{Cr}(\mathrm{VI})$ or nickel. Similar results were also observed across smoking group strata. The joint effect of smoking and $\mathrm{Cr}(\mathrm{VI})$ or nickel generally exceeded additivity. Various sensitivity analyses corroborated the robustness of these results. Although differential bias in our results due to combined Berkson and classical error structure cannot be ruled out, our results warrant a continuing awareness to monitor the impact of occupational metal exposure on human cancer by epidemiologic, toxicological, and experimental investigations.

## AUTHOR CONTRIBUTIONS

KS, T. Brüning, HK, and RV contributed to the original conception and acquired funding for the project. KS, J. Schüz, HK, RV, AO, BP, T. Brüning, and T. Behrens developed the methodology for this analysis. T. Behrens. wrote the draft, by integrating suggestions from CG, AO, SP, J. Schüz, KS, HK, RV, BP, and J. Siemiatycki.

CG, BK, and LP conducted the formal statistical analyses. J. Siemiatycki, M-EP, P. Gustavsson, PB, P. Guenel, DL, SK, H-EW, MTL, NEC, D. Mirabelli, LR, DC, LS, K-HJ, WA, HP, GF-T, DZ, JKF, JL, BS, JRMcL, PAD, VB, LF, VJ, TP, EF, D. Mates, FF, and BBdM participated in data acquisition, design, and quality assurance of the original study data. Data curation of the pooled data (annotation, data cleaning, and maintaining research data) was done by AO, BK, SP, and LP. All authors participated in critical revision of the manuscript and provided approval of the finalized submitted version.

The work reported in the paper has been performed by the authors, unless clearly specified in the text.

## CONFLICT OF INTEREST

None reported.

## DATA AVAILABILITY STATEMENT

The data are from a pooled study, and the data right owners are the individual study groups whose authors may be contacted at synergy@iarc.fr. Further information is available from the corresponding author upon request.

## DISCLAIMER

T. Behrens, BK, BP, and T. Brüning, as staff of the Institute for Prevention and Occupational Medicine (IPA), are or were formerly employed by the study's main financing body, the German Social Accident Insurance. IPA is an independent research institute of the Ruhr University Bochum. The authors alone are responsible for the views expressed in this article and they do not necessarily represent the decisions, policy or views of their affiliated institutes.

## ETHICS STATEMENT

The IARC Institutional Review Board provided ethical approval for the pooled study, while national ethics committees approved the local case-control studies. All study subjects gave written informed consent.

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## FIGURE LEGENDS

## Figure 1:

Exposure-response relationships among males and females for cumulative hexavalent chromium and nickel exposure with different lag periods, adjusted for study, age group, cigarette pack-years, time-since-quitting smoking, and ever employment in a 'list A' job. The histograms on the x-axis show the distribution of the cumulative exposure in the respective sub-populations.

## TABLES:

Table 1. Descriptive characteristics of the study participants (16,901 lung-cancer cases, 20,965 control subjects) by exposure to hexavalent chromium $(\mathrm{Cr}(\mathrm{VI}))$ and nickel

| Characteristic | Exposure category | Exposed to Ni or $\mathrm{Cr}(\mathrm{VI})$ |  |  |  | Unexposed to Ni and $\mathrm{Cr}(\mathrm{VI})$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Cases |  | Controls |  | Cases |  | Controls |  |
|  |  | No. (\%) | Median (IQR) | No. (\%) | Median (IQR) | No. (\%) | Median (IQR) | No. (\%) | Median (IQR) |
| MEN |  | 4,135 |  | 3,823 |  | 9,470 |  | 12,628 |  |
|  | Median (IQR) |  | 63 (13) |  | 63 (13) |  | 64 (12) |  | 63 (13) |
| Age [years] | <45 | 132 (3) |  | 177 (5) |  | 354 (4) |  | 718 (6) |  |
|  | 45-65 | 2,260 (55) |  | 1,955 (51) |  | 4,725 (50) |  | 6,227 (49) |  |
|  | 65+ | 1,743 (42) |  | 1,691 (44) |  | 4,391 (46) |  | 5,683 (45) |  |
| Smoking status | Never smoker | 116 (3) |  | 846 (22) |  | 374 (4) |  | 3,591 (28) |  |
|  | Former smoker | 1,397 (34) |  | 1,789 (47) |  | 3,390 (36) |  | 5,539 (44 |  |
|  | Current smoker | 2,622 (63) |  | 1,188 (31) |  | 5,706 (60) |  | 3,498 (28) |  |
| Cigarette pack-years (current and former smokers) | $<10$ | 202 (5) |  | 594 (20) |  | 490 (5) |  | 2,130 (24) |  |
|  | 10-<20 | $411(10)$ $1533(39)$ |  | 586(20) |  | $\begin{array}{r}837 \\ \hline \text { (9) }\end{array}$ |  | 1,746 (19) |  |
|  | 20-<40 | 1,533 (39) |  | 1,067 (36) |  | 3,336 (37) |  | 3,004 (33) |  |
|  | 40+ | 1,873 (47) |  | 730 (24) |  | 4,433 (49) |  | 2,157 (24) |  |
| Years-since-quitting smoking (former smokers) | >2-7 | 521 (37) |  | 306 (17) |  | 1,225 (36) |  | 912 (16) |  |
|  | 8-15 | 394 (28) |  | 429 (24) |  | 961 (28) |  | 1,262 (23) |  |
|  | 16-25 | 297 (21) |  | 516 (29) |  | 747 (22) |  | 1,579 (29) |  |
|  | >25 | 185 (13) |  | 538 (30) |  | 457 (13) |  | 1,786 (32) |  |
| Employed in 'List A' job | Ever | 922 (22) |  | 668 (17) |  | 807 (9) |  | 656 (5) |  |
|  | Adenocarcinoma | 896 (22) |  |  |  | 2,429 (26) |  |  |  |
| Lung-cancer cell type | Squamous cell carcinoma | 1,885 (46) |  |  |  | 3,943 (42) |  |  |  |
|  | Small-cell lung cancer | 703 (17) |  |  |  | 1,497 (16) |  |  |  |
|  | Other/unspecified | 625 (15) |  |  |  | 1,548 (16) |  |  |  |
|  | Not available | 26 (1) |  |  |  | 53 (1) |  |  |  |
| Nickel [ $\mu \mathrm{g} / \mathrm{m}^{3}$-years] |  |  | 22.7 (64) |  | 21.5 (60) |  |  |  |  |
| $\mathrm{Cr}(\mathrm{VI})$ [ $\mu \mathrm{g} / \mathrm{m}^{3}$-years] |  |  | 42.8 (91) |  | 40.8 (86) |  |  |  |  |


| Characteristic | Exposure category | Exposed to Ni or $\mathrm{Cr}(\mathrm{VI})$ |  |  |  | Unexposed to Ni and $\mathrm{Cr}(\mathrm{VI})$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Cases |  | Controls |  | Cases |  | Controls |  |
|  |  | No. (\%) | Median (IQR) | No. (\%) | Median (IQR) | No. (\%) | Median (IQR) | No. (\%) | Median (IQR) |
| WOMEN |  | 161 |  | 146 |  | 3,135 |  | 4,368 |  |
| Age [years] | Median (IQR) |  | 63 (14) |  | 62 (15) |  | 61 (16) |  | 61 (17) |
| Smoking status | <45 | 11 (7) |  | 5 (3) |  | 218 (7) |  | 471 (11) |  |
|  | 45-64 | 82 (51) |  | 75 (51) |  | 1,696 (54) |  | 2,097 (48) |  |
|  | 65+ | 68 (42) |  | 66 (45) |  | 1,221 (39) |  | 1,800 (41) |  |
|  | Never smoker | 35 (22) |  | 76 (52) |  | 844 (27) |  | 2,640 (60) |  |
|  | Former smoker | 24 (15) |  | 35 (24) |  | 621 (20) |  | 857 (20) |  |
|  | Current smoker | 102 (63) |  | 35 (24) |  | 1,670 (53) |  | 871 (20) |  |
| Cigarette pack-years (current and former smokers) | $<10$ | $9 \text { (7) }$ |  | 20 (29) |  | 222 (10) |  | 629 (36) |  |
|  | 10-19 | 19 (15) |  | 17 (24) |  | 377 (16) |  | 403 (23) |  |
|  | 20-<40 | 53 (42) |  | 25 (36) |  | 906 (40) |  | 464 (27) |  |
|  | 40+ | 45 (36) |  | 8 (11) |  | 786 (34) |  | 232 (13) |  |
| Years-since-quitting smoking (former smokers) | 2-7 years | 9 (38) |  | 7 (20) |  | 271 (44) |  | 197 (23) |  |
|  | $8-15$ years | 6 (25) |  | 5 (14) |  | 170 (27) |  | 202 (24) |  |
|  | 16-25 years | 6 (25) |  | 13 (37) |  | 121 (19) |  | 238 (28) |  |
|  | $26+$ years | 3 (12) |  | 10 (29) |  | 59 (10) |  | 220 (26) |  |
| Employed in 'List A' job | Ever | 17 (11) |  | 16 (11) |  | 41 (1) |  | 24 (1) |  |
| Lung-cancer cell type | Not available | 1 (1) |  |  |  | 14 (1) |  |  |  |
|  | Adenocarcinoma | 56 (35) |  |  |  | 1,371 (44) |  |  |  |
|  | Squamous cell carcinoma | 37 (23) |  |  |  | 638 (20) |  |  |  |
|  | Small-cell lung cancer | 37 (23) |  |  |  | 493 (16) |  |  |  |
|  | Other/unspecified | 30 (19) |  |  |  | 619 (20) |  |  |  |
| Nickel [ $\mu \mathrm{g} / \mathrm{m}^{3}$-years] | Median (IQR) |  | 16.7 (30) |  | 14.2 (29) |  |  |  |  |
| $\mathrm{Cr}(\mathrm{VI})\left[\mu \mathrm{g} / \mathrm{m}^{3}\right.$-years] | Median (IQR) |  | 26.2 (46) |  | 26.0 (46) |  |  |  |  |

IQR = Interquartile range
$\mathrm{Cr}(\mathrm{VI})=$ Hexavalent chromium

Table 2. Lung cancer odds ratios (OR) and $95 \% \mathrm{Cl}$ in relation to indices of occupational exposure to nickel and hexavalent chromium in the SYNERGY Study

| Indices of occupational exposure | Exposure category | Men |  |  |  |  |  |  | Women |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Cases | Contr. | OR1 | 95\%CI | OR2 | 95\%CI | 99.4\%CI ${ }^{\circ}$ | Cases | Contr. | OR1 | 95\%CI | OR2 | 95\%CI | 99.4\%CI ${ }^{\circ}$ |
| Nickel | Never | 10,389 | 13,311 | 1.0 | Ref. | 1.0 | Ref. | Ref. | 3,145 | 4,383 | 1.0 | Ref. | 1.0 | Ref. | Ref. |
|  | Ever | 3,216 | 3,140 | 1.27 | 1.20-1.35 | 1.12 | 1.05-1.20 | 1.02-1.23 | 151 | 131 | 1.64 | 1.28-2.10 | 1.23 | 0.93-1.63 | 0.83-1.83 |
|  | 1-9 | 1,273 | 1,375 | 1.13 | 1.04-1.23 | 1.01 | 0.92-1.11 | 0.88-1.15 | 102 | 79 | 1.84 | 1.36-2.50 | 1.36 | 0.96-1.91 | 0.84-2.20 |
| Duration (years) | 10-19 | 609 | 575 | 1.35 | 1.20-1.52 | 1.15 | 1.00-1.31 | 0.95-1.39 | 30 | 36 | 1.25 | 0.76-2.06 | 0.95 | 0.54-1.66 | 0.43-2.11 |
|  | 20-29 | 487 | 419 | 1.47 | 1.29-1.68 | 1.26 | 1.08-1.46 | 1.01-1.56 | 13 | 8 | 2.02 | 0.84-5.15 | 1.53 | 0.58-4.32 | 0.38-6.26 |
|  | 30+ | 847 | 771 | 1.37 | 1.23-1.51 | 1.23 | 1.09-1.38 | 1.04-1.44 | 6 | 8 | 1.03 | 0.33-2.99 | 0.87 | 0.25-2.85 | 0.16-4.74 |
| Test for trend, $p$-value § |  |  |  |  | <0.001 |  | <0.001 |  |  |  |  | 0.05 |  | 0.59 |  |
| Excl. never exposed |  |  |  |  | <0.001 |  | 0.001 |  |  |  |  | 0.51 |  | 0.87 |  |
| $\mathrm{Cr}(\mathrm{VI})$ | Never | 9,474 | 12,631 | 1.0 | Ref. | 1.0 | Ref. | Ref. | 3,137 | 4,368 | 1.0 | Ref. | 1.0 | Ref. | Ref. |
|  | Ever | 4,131 | 3,820 | 1.38 | 1.31-1.46 | 1.21 | 1.14-1.28 | 1.11-1.31 | 159 | 146 | 1.52 | 1.20-1.92 | 1.14 | 0.87-1.49 | 0.78-1.67 |
| Duration (years) | 1-9 | 1,518 | 1,628 | 1.17 | 1.09-1.27 | 1.04 | 0.95-1.13 | 0.92-1.18 | 110 | 90 | 1.69 | 1.27-2.26 | 1.28 | 0.93-1.77 | 0.81-2.03 |
|  | 10-19 | 780 | 672 | 1.52 | 1.36-1.69 | 1.27 | 1.13-1.44 | 1.07-1.51 | 30 | 38 | 1.17 | 0.71-1.91 | 0.86 | 0.49-1.49 | 0.39-1.89 |
|  | 20-29 | 647 | 531 | 1.60 | 1.42-1.80 | 1.34 | 1.17-1.53 | 1.10-1.62 | 13 | 9 | 1.77 | 0.76-4.34 | 1.24 | 0.48-3.34 | 0.32-4.81 |
| Test for trend, $p$-value § | 30+ | 1,186 | 989 | 1.53 | 1.40-1.67 | 1.37 | 1.23-1.51 | 1.18-1.58 | 6 | 9 | 0.89 | 0.30-2.51 | 0.75 | 0.22-2.35 | 0.14-3.88 |
|  |  |  |  |  | <0.001 |  | <0.001 |  |  |  |  | 0.12 |  | 0.99 |  |
| Excl. never exposed |  |  |  |  | <0.001 |  | <0.001 |  |  |  |  | 0.58 |  | 0.79 |  |
| Ni or $\mathrm{Cr}(\mathrm{VI})$ | Never | 9,470 | 12,628 | 1.0 | Ref. | 1.0 | Ref. | Ref. | 3,135 | 4,368 | 1.0 | Ref. | 1.0 | Ref. | Ref. |
|  | Ever | 4,135 | 3,823 | 1.39 | 1.31-1.46 | 1.21 | 1.14-1.28 | 1.11-1.31 | 161 | 146 | 1.54 | 1.22-1.95 | 1.15 | 0.88-1.51 | 0.79-1.68 |
| Ni and $\mathrm{Cr}(\mathrm{VI})$ | Never | 10,393 | 13,314 | 1.0 | Ref. | 1.0 | Ref. | Ref. | 3,147 | 4,383 | 1.0 | Ref. | 1.0 | Ref. | Ref. |
|  | Ever | 3,212 | 3,173 | 1.27 | 1.20-1.35 | 1.12 | 1.05-1.20 | 1.02-1.23 | 149 | 131 | 1.62 | 1.23-2.07 | 1.22 | 0.92-1.61 | 0.82-1.81 |
| Cumulative exposure to nickel [ $\mu \mathrm{g} / \mathrm{m}^{3}$-years] | $>0-\leq 11.9$ | 672 | 771 | 1.04 | 0.93-1.16 | 0.92 | 0.81-1.04 | 0.77-1.09 | 54 | 47 | 1.70 | 1.14-2.56 | 1.29 | 0.82-2.05 | 0.68-2.47 |
|  | >11.9- $\leq 30.9$ | 749 | 775 | 1.18 | 1.06-1.31 | 1.06 | 0.94-1.20 | 0.89-1.26 | 47 | 42 | 1.61 | 1.05-2.48 | 1.34 | 0.83-2.16 | 0.68-2.63 |
|  | >30.9- $\leq 78.1$ | 914 | 790 | 1.44 | 1.30-1.59 | 1.20 | 1.07-1.35 | 1.02-1.41 | 29 | 28 | 1.45 | 0.85-2.48 | 0.97 | 0.53-1.76 | 0.41-2.25 |
|  | >78.1 | 881 | 804 | 1.42 | 1.29-1.57 | 1.29 | 1.15-1.45 | 1.10-1.52 | 21 | 14 | 1.91 | 0.97-3.90 | 1.29 | 0.60-2.86 | 0.43-3.87 |



OR 1 is adjusted for study and age group
OR2 is adjusted for study, age group, smoking (log(cigarette pack-years+1), time-since-quitting smoking (current smokers, stopping smoking 2-7 years, 8-15 years, 16-25,
26+ years before interview/diagnosis, never smokers)), and List A jobs

* OR2 in "time since last exposure" is in addition adjusted for duration (continuous) of exposure
${ }^{\circ} 99.4 \%$ CI Bonferroni-corrected for 9 subtests

| Cumulative exposure [ $\mu \mathrm{g} / \mathrm{m}^{3}$-years] | Population-based studies |  |  |  | Hospital-based studies |  |  |  | Blue-collar workers only |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cases/Ctrls. | OR2 | 95\%CI | 99.4\%CI ${ }^{\circ}$ | Cases/Ctrls. | OR2 | $295 \% \mathrm{Cl}$ | 99.4\% $\mathrm{Cl}^{\circ}$ | Cases/Ctrls. | OR2 | 95\%CI | 99.4\% $\mathrm{Cl}^{\circ}$ |
| Unexposed | 6,916/9,815 | 1.0 | Ref. | Ref. | 2,478/2,685 | 1.0 | Ref. | Ref. | 6,773/7,518 | 1.0 | Ref. | Ref. |
| $>0-\leq 15.3$ | 740/707 | 1.17 | 1.03-1.32 | 0.97-1.36 | 150/139 | 1.09 | 0.83-1.42 | 0.74-1.49 | 963/938 | 1.02 | 0.92-1.14 | 0.87-1.20 |
| $>15.3-\leq 40.3$ | 774/685 | 1.29 | 1.14-1.46 | 1.07-1.50 | 202/205 | 0.97 | 0.77-1.21 | 0.72-1.33 | 1028/945 | 1.08 | 0.97-1.20 | 0.93-1.26 |
| >40.3-599.5 | 724/671 | 1.27 | 1.12-1.44 | 1.05-1.49 | 263/223 | 1.06 | 0.86-1.31 | 0.78-1.37 | 1016/947 | 1.09 | 0.97-1.21 | 0.93-1.26 |
| >99.5 | 775/700 | 1.36 | 1.20-1.54 | 1.18-1.70 | 378/308 | 1.16 | 0.97-1.39 | 0.91-1.50 | 1117/965 | 1.21 | 1.09-1.35 | 1.04-1.41 |
| Test for trend, $p$-value | <0.001 |  |  |  | 0.32 |  |  |  |  |  | 0.01 |  |
| Excl. never exposed | 0.15 |  |  |  | 0.59 |  |  |  | 0.07 |  |  |  |
| Cumulative exposure [ $\mu \mathrm{g} / \mathrm{m}^{3}$-years] | Restricted to workers starting iobs |  |  |  | Restricted to workers starting iobs |  |  |  | Excluding regular welders |  |  |  |
|  | 1960 or later |  |  |  | 1970 or later |  |  |  |  |  |  |  |
|  | Cases/Ctrls. | OR2 | 95\%CI | 99.4\%Cl ${ }^{\circ}$ | Cases/Ctris. | OR2 | 95\%CI | 99.4\%CI ${ }^{\circ}$ | Cases/Ctrls. OR2 |  | 95\%CI | 99.4\%C1 ${ }^{\circ}$ |
| Unexposed | 2,379/3,651 | 1.0 | Ref. | Ref. | 758/1,539 | 1.0 | Ref. | Ref. | 9,474/12,631 | 1.0 | Ref. | Ref. |
| $>0-\leq 15.3$ | 198/216 | 1.08 | 0.86-1.37 | 0.78-1.51 | 56/77 | 1.110. | 0.74-1.68 | 0.62-2.00 | 936/921 | 1.10 | 0.99-1.22 | 0.95-1.30 |
| $>15.3-\leq 40.3$ | 224/234 | 1.26 | 1.01-1.58 | 0.92-1.74 | 67/90 | 1.310. | 0.89-1.92 | 0.76-2.24 | 954/883 | 1.18 | 1.06-1.31 | 1.02-1.39 |
| >40.3-<99.5 | 186/223 | 1.04 | 0.82-1.32 | 0.74-1.45 | 48/83 | 0.86 | 0.57-1.31 | 0.48-1.56 | 879/846 | 1.18 | 1.05-1.31 | 1.00-1.37 |
| 9.5 | 212/207 | 1.30 | 1.03-1.64 | 0.93-1.81 | 36/42 | 1.280. | 0.76-2.17 | 0.61-2.69 | 843/785 | 1.25 | 1.12-1.40 | 1.07-1.48 |
| Test for trend, p-value | 0.03 |  |  |  | 0.79 |  |  |  | <0.001 |  |  |  |
| Excl. never exposed | 0.30 |  |  |  | 0.35 |  |  |  | 0.07 |  |  |  |

Table 3b. Lung cancer odds ratios (OR) and $95 \% \mathrm{Cl}$ in relation to cumulative exposure to nickel in subgroups of men in the SYNERGY Study

| Cumulative exposure [ $\mu \mathrm{g} / \mathrm{m}^{3}$-years] | Population-based studies |  |  |  | Hospital-based studies |  |  |  | Blue-collar workers only |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cases/Ctrls. | OR2 | 95\%CI | 99.4\%Cl ${ }^{\circ}$ | Cases/Ctrls. | OR2 | 95\%CI | 99.4\%CI ${ }^{\circ}$ | Cases/Ctrls. | OR2 | 95\%CI | 99.4\%Cl ${ }^{\circ}$ |
| Unexposed | 7,669/10,359 | 1.0 | Ref. | Ref. | 2,657/2,834 | 1.0 | Ref. | Ref. | 7,685/8,195 | 1.0 | Ref. | Ref. |
| $>0-\leq 11.9$ | 431/484 | 0.94 | 0.81-1.10 | 0.78-1.16 | 138/149 | 0.87 | 0.67-1.14 | 0.59-1.19 | 672/769 | 0.84 | 0.75-0.96 | 0.71-1.01 |
| >11.9- 30.9 | 526/525 | 1.14 | 0.98-1.32 | 0.92-1.36 | 174/177 | 0.92 | 0.72-1.17 | 0.67-1.28 | 749/770 | 0.97 | 0.86-1.09 | 0.82-1.15 |
| >30.9-¢78.1 | 650/583 | 1.20 | 1.05-1.37 | 1.01-1.47 | 240/189 | 1.17 | 0.94-1.46 | 0.86-1.57 | 912/785 | 1.10 | 0.98-1.23 | 0.93-1.29 |
| >78.1 | 653/627 | 1.29 | 1.13-1.48 | 1.06-1.56 | 262/211 | 1.22 | 0.99-1.51 | 0.98-1.76 | 879/794 | 1.18 | 1.05-1.33 | 1.00-1.39 |
| Test for trend, p-value |  |  | 0.004 |  |  |  | 0.06 |  |  |  | 0.03 |  |
| Excl. never exposed |  |  | 0.12 |  |  |  | 0.03 |  |  |  | 0.01 |  |

Table 4. Lung-cancer odds ratios and $95 \% \mathrm{CI}$, p-value for multiplicative interaction and relative excess risk due to interaction (RERI) and 95\% Cl in relation to occupational chromium ( VI ) and nickel exposure and smoking among men and women

| Exposure status | Men |  |  |  |  | Women |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Cases | Controls | OR* | 95\%CI | 99.4\%CI ${ }^{\circ}$ | Cases | Controls | OR* | 95\%CI | 99.4\% $\mathrm{Cl}^{\circ}$ |
| Chromium (VI) |  |  |  |  |  |  |  |  |  |  |
| Never smoker and never $\mathrm{Cr}(\mathrm{VI})$ | 374 | 3,592 | 1.0 | Ref. | Ref. | 844 | 2,640 | 1.0 | Ref. | Ref. |
| Never smoker and $\mathrm{Cr}(\mathrm{VI})$ | 116 | 845 | 1.22 | 0.98-1.53 | 0.89-1.68 | 35 | 76 | 1.13 | 0.73-1.73 | 0.62-2.08 |
| Ever smoker and never $\mathrm{Cr}(\mathrm{VI})$ | 9,100 | 9,039 | 9.31 | 8.34-10.42 | 7.95-10.89 | 2,293 | 1,728 | 4.77 | 4.29-5.31 | 4.10-5.55 |
| Ever smoker and ever $\mathrm{Cr}(\mathrm{VI})$ | 4,015 | 2,975 | 11.63 | 10.34-13.11 | 9.83-13.75 | 124 | 70 | 6.22 | 4.54-8.57 | 3.97-9.73 |
| $p$-value multiplicative interaction |  |  | 0.86 |  |  |  |  | 0.60 |  |  |
| RERI with linear model** |  |  | 2.10 | 1.41-2.79 |  |  |  | 1.31 | -0.66-3.29 |  |
| Nickel |  |  |  |  |  |  |  |  |  |  |
| Never smoker and never nickel | 402 | 3,735 | 1.0 | Ref. | Ref. | 847 | 2,650 | 1.0 | Ref. | Ref. |
| Never smoker and nickel | 88 | 702 | 1.08 | 0.84-1.38 | 0.76-1.53 | 32 | 66 | 1.27 | 0.80-1.97 | 0.67-2.39 |
| Ever smoker and never nickel | 9,987 | 9,576 | 9.31 | 8.37-10.38 | 7.99-10.84 | 2,298 | 1,733 | 4.77 | 4.29-5.31 | 4.10-5.55 |
| Ever smoker and ever nickel | 3,128 | 2,438 | 10.63 | 9.46-11.98 | 9.00-12.56 | 119 | 65 | 6.60 | 4.78-9.18 | 4.16-10.46 |
| $p$-value multiplicative interaction |  |  | 0.66 |  |  |  |  | 0.76 |  |  |
| RERI with linear model** |  |  | 1.25 | 0.56-1.93 |  |  |  | 1.56 | -0.60-3.72 |  |

* OR adjusted for study, age group and "List A" jobs
** Confidence intervals are based on 1,000 bootstrap samples
- $99.4 \% \mathrm{Cl}$ Bonferroni-corrected for 9 subtests


Cumulative chromium VI exposure ( $\mu \mathrm{g} / \mathrm{m}^{3}$-yrs)
Males



Females



- Assessment of lung-cancer risk after occupational exposure to $\mathrm{Cr}(\mathrm{VI})$ and nickel in a pooled case-control study to investigate quantitative exposure-effect relationships using secondary measurement data.
- Cubic-spline analyses showed linear exposure-effect relationship for nickel and $\mathrm{Cr}(\mathrm{VI})$ among women, and also for $\mathrm{Cr}(\mathrm{VI})$ among men. The exposure-effect relationship for nickel among men was sublinear at lower cumulative exposure levels.
- Exploring various lagging periods did not change results.
- Various sensitivity analyses corroborated the robustness of these results.

