

# Ecological network complexity scales with area

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#### 1 Ecological network complexity scales with area

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57 Larger geographical areas contain more species –an observation raised to a law in ecology. 58 Less explored is whether biodiversity changes are accompanied by a modification of 59 interaction networks. We use data from 32 spatial interaction networks from different 60 ecosystems to analyse how network structure changes with area. We find that basic 61 community structure descriptors (number of species, links, and links per species) increase 62 with area following a power-law. Yet, the distribution of links per species varies little with 63 area, indicating that the fundamental organization of interactions within networks is 64 conserved. Our null model analyses suggest that the spatial scaling of network structure is 65 determined by factors beyond species richness and the number of links. We demonstrate 66 that biodiversity-area relationships can be extended from species counts to higher levels of 67 network complexity. Therefore, the consequences of anthropogenic habitat destruction 68 may extend from species loss to wider simplification of natural communities.

69

#### 70 INTRODUCTION

# 71 Biotic interactions, the missing link

72 Research on the spatial scaling of biodiversity has historically focused on the increase of species richness with area<sup>1-3</sup> and on other components of biodiversity, such as functional or 73 74 phylogenetic diversity<sup>4-6</sup>. The Species Area Relationship (SAR) is essential to estimate species 75 richness in a region and species loss following habitat destruction and range contraction<sup>7–9</sup>. 76 Ecological communities are, however, more than disconnected collections of species. Instead, 77 they can be represented as networks, with species as nodes and interactions among species as 78 links<sup>10</sup>. Ecological interactions are fundamental to predict ecosystem responses to environmental changes<sup>11-13</sup> and to sustain important ecosystem functions<sup>14-16</sup>. Understanding how the structure 79 80 of interaction networks changes with area is, thus, crucial to fully characterize the spatial scaling 81 of biodiversity and to predict ecosystem responses to human activities.

82 The influence of spatial processes on the organization of interaction networks has long
 83 interested ecologists<sup>17–19</sup>. However, research on the spatial scaling of network structure has been

scarce<sup>20-22</sup>. This scaling concerns two hierarchical levels: the number of building blocks within 84 85 communities (species and their interactions) and the relationships between them. The scaling of the number of links (i.e., biotic interactions) with area has been previously predicted<sup>20</sup> by 86 87 unifying the SAR with the well-established scaling of the number of links with species richness<sup>23–26</sup>. Similarly, recent theoretical research has shown that a number of network-area 88 89 relationships (NARs) can emerge from different spatial mechanisms and assembly processes, 90 such as different SARs across trophic levels or dispersal limitation<sup>21</sup>. Yet, empirically 91 documenting the specific shape of these relationships and assessing their potential universality 92 across biomes, interaction types, and spatial domains remains a major challenge. 93 The power function of the form  $S \sim cA^z$  has been found to describe the increase in species richness with area A across all ecosystem types<sup>1,27</sup>, with parameter z varying 94 substantially with the spatial extent studied<sup>27–29</sup>. Whether other aspects of network complexity 95 can be captured by the same functional form, and how the parameters of such scaling vary with 96 97 network properties and spatial extent, are questions of fundamental importance for 98 characterizing the effects of area loss on biodiversity. Similarly, as many aspects of network 99 structure can change with species diversity and network connectance<sup>30–33</sup> (i.e., the proportion of 100 realised interactions among all potential ones), it is important to determine whether the spatial 101 scaling of network structure is a trivial consequence of the increase of species richness with 102 area, or whether such scaling of network properties follow rules beyond those applying to 103 species richness.

104

# A unique dataset across ecosystems and biomes

We compiled 32 datasets from different ecosystems across the globe describing interaction networks (including both mutualistic and antagonistic interactions) within two spatial domains: regional and biogeographical. The regional domain represents communities for which sampling was conducted locally in a replicated fashion within narrow spatial extents (i.e., maximum spatial extent of ca. 1000 km<sup>2</sup>). The biogeographical spatial domain includes datasets for which the sampling units span much broader areas, encompassing multiple biomes (see

111 Methods and Supplementary Text 1). Thus, biogeographical data comprise communities 112 exposed to larger environmental heterogeneity, stronger dispersal barriers and historical 113 contingencies, which combine to produce diversity patterns at large spatial scales. Hence, we 114 expected *a priori* the scaling relationships for the biogeographical domain to be steeper than 115 those for the regional domain, as within the biogeographical domain, species assemblages can 116 be evolutionary less related and exhibit greater turnover of species and links. Each dataset 117 contained interaction data, using different methodologies to document pairwise interactions. To 118 characterise changes in network properties with changes in area, we sequentially aggregated the 119 sampling units available, scoring the structure of the network at each step of the aggregation 120 procedure (see Methods).

# 121 RESULTS & DISCUSSION

### 122 Links or species: which one is gained faster?

123 We found that network complexity increased with area at all levels: at the level of the number of 124 building blocks (species and links) and at the level of their combinations (links per species; 125 Figure 1; Table 1 and Supplementary Table 1). For all datasets, the relationships followed a 126 power-law function, regardless of the spatial extent of observation or interaction type. Within 127 both the regional and biogeographical domain, we found that an extended power function best 128 describes the scaling of network complexity with area (see Methods and Supplementary Table 1). This function has the form  $N = cA^{(zA^{-d})}$ , where N is a given network property, A is area, 129 130 and c, z and d are fitted parameters, where z represents the slope of a given NAR in log-log 131 space (i.e., the scaling exponent) and *d* controls the strength of the asymptotic flattening. 132 We found systematic differences in the shape of the scaling relationships between the 133 regional and biogeographical domains. All measures of network complexity followed a linear-134 concave increase with area size in the regional domain (i.e.  $z \gg d > 0$ ) while in the 135 biogeographical domain, the increase was convex for most datasets (z > 0 > d) (but see 136 Supplementary Figure 7 and Supplementary Figure 8). The biogeographic domain also showed

137 larger variability across datasets than the regional domain, suggesting lower predictability at138 larger spatial extents (Supplementary Table 3).

139 The number of links increased faster with area than the number of species, within both 140 the regional and biogeographical domains. Importantly, whether links increase faster than 141 species with area will depend on how the number of links scales with species richness<sup>20,21</sup>. Previous empirical studies<sup>23-26</sup> have used a power function to relate the increase of the number 142 143 of links with species richness, and found the value of the scaling exponents to lie between 1 and  $2^{25,26}$  – coining the expressions of the "link-species scaling law" <sup>34</sup> versus "constant connectance" 144 145 hypothesis<sup>24</sup>, respectively. Consistent with previous work, we found the increase in the number 146 of links with an increase in species richness to follow a power law in all data sets (Figure 2, 147 Supplementary Table 2). The exponents were larger for the biogeographical (mean  $\pm$  SD: z =148  $1.78 \pm 0.20$ ) than for the regional domain ( $z = 1.60 \pm 0.20$ ). For both spatial domains, the 149 scaling exponents of the links-per-species relationship were thus clustered between 1.5 and 2, 150 but with substantial variability in specific values (Figure 2, Supplementary Table 2). The high 151 variation observed suggests that the species richness of a community may carry little 152 information on how the number of links will change with area. Instead, the scaling of the

153 number of links with species richness may need to be established on a network-specific basis<sup>20</sup>.

# 154 The unpredictability of vertical diversity

155 Within a network, the links can be organized in multiple ways. In our dataset, the mean number 156 of resources used by a consumer (i.e. mean indegree: Links/Species<sub>consumer</sub>) increased with area. 157 The general shape of the relationship was similar to that of the other complexity measures 158 analysed, i.e. linear-concave in the regional domain and convex in the biogeographical domain 159 (Figure 3, Supplementary Table 1). Still, compared to the relationships observed for other 160 network properties, we observed greater variability in indegree-area relationships within spatial 161 domains (Supplementary Table 3). This variability is partially driven by the variability in 162 changes in consumer to resource ratios with area observed across datasets (Supplementary Table 163 1). Although previous studies have shown a consistent increase of the slope of SARs with

164 trophic level<sup>35,36</sup>, we did not observe any consistent increase nor decrease of the

165 consumer: resource richness ratio with area at any of the spatial domains (Supplementary Table

- 166 1). These findings suggest that the spatial scaling of vertical diversity properties, such as the
- 167 proportion of species per trophic level, might be system-specific.

# 168 The scale-invariance of degree distributions

169 Bevond the network properties considered above, other network properties were more consistent 170 across scales. The degree distribution of ecological networks is typically highly skewed, with many specialist species and few generalists<sup>37,38</sup>. This skewness may influence the stability and 171 172 robustness of communities<sup>38</sup> and it is tightly linked to widely studied network properties, such as nestedness<sup>39,40</sup>, (i.e. the degree to which specialist species interact with subsets of the species 173 174 interacting with generalists). Previous theoretical work has suggested that the skewness of ecological networks is preserved across spatial scales<sup>21</sup>, an expectation borne out in our data. 175 176 Despite variability in the degree distribution across datasets, the same function provided a good 177 fit to the degree distribution of most ecological networks across the full range of areas (see 178 Methods and Figure 3). Although we observed variation in the parameters of the functions 179 (Supplementary Figure 4), the consistency in the fundamental shape suggests that community 180 robustness to species loss (independently of whether it is high or low) may be maintained across 181 spatial scales.

# 182 Disentangling the effect of species richness

183 Given the influence of species richness and network connectance on many aspects of network structure<sup>30–33</sup>, we aimed to investigate whether network structure changes with area beyond 184 185 those changes associated to the increase in species richness and links with area. To do so, we 186 generated random networks with two different null models, as reflecting the case where the 187 spatial scaling of network structure emanates from the spatial scaling of species richness alone 188 (with no change in links per species), or the spatial scaling of both species richness and links, 189 respectively. In both models, we started from the metaweb (i.e. the overall network pooled 190 across all original networks within each specific study), then randomly drew the same number

191 of species as observed in the original, local networks. In null model 1, local random networks 192 were drawn as random subsets of the metaweb, picked to have the same species richness as the 193 observed local networks. In other words, after randomly selecting a given number of species 194 from the metaweb, we only kept interactions among this subset of species, thereby determining 195 the number of links of the network generated. Thus, this procedure mimics a scenario where 196 random local networks have the same number of species as the observed local networks, but the 197 number of links and the associated network properties may differ. In null model 2, we preserved 198 both species richness and the number of interactions observed in the original networks, but 199 distributed the links randomly among the species (see Methods). Accordingly, null model 2 200 mimics a scenario where a change in area causes a potential change in indegree and network 201 degree distribution, while not affecting either the number of species or links. For each of the 202 100 randomizations under each scenario, we then scored each of the complexity metrics defined 203 above: species, links, links per species, indegree and network degree distribution.

204 The observed spatial scaling of network structure in the original networks differed 205 significantly from the results generated by the null models. The number of links in empirical 206 networks increased more slowly with an increase in the number of species than expected from a 207 random sample of species from the metaweb (i.e. null model 1) (Supplementary Figure 5). This 208 slower increase is driven by the fact that local original networks show greater complexity (i.e. 209 more links per species) than random networks. This result suggests that the scaling of the 210 number of links is shaped by factors beyond the species richness of a community, such as 211 evolutionary constraints, phenological matching or competition. In contrast, other structural 212 properties, such as mean indegree and degree distributions, are consistent with the pattern 213 derived from null model 1 (Supplementary Figure 5 and Supplementary Figure 6; 214 Supplementary Table 5 and Supplementary Table 6). The latter consistency indicates that these 215 particular network properties might be inherited from the metaweb, and that the particular co-216 occurrence structure of species in the empirical data is not meaningfully impacting the degree 217 distribution. This interpretation is supported by patterns observed under null model 2. When

links are reshuffled randomly amongst selected species from the metaweb, the pattern is
substantially different from the observed one: here, both the mean indegree and network degree
distributions of the randomized networks strongly deviate from that of their original
counterparts (Supplementary Figure 5 and Supplementary Figure 6; Supplementary Table 5 and
Supplementary Table 6). Similarly, the observed spatial scaling of the consumer:resource ratio
was poorly predicted by both null models.

- Overall, our null model analyses suggest that the spatial scaling of network structure is determined by factors beyond species richness and the number of links. With a change in area, we see changes in important features of ecological networks – such as vertical diversity, consumer specialisation, and degree distribution – well beyond those expected if changes were driven by the number of species and/or links alone.
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### The fragility of biotic interactions

230 In conclusion, our analyses of multiple interaction networks from different biomes, 231 interaction types, and spatial domains allowed us to explore the generality in the spatial scaling 232 of several structural properties of ecological networks. Our results demonstrate how previouslyestablished biodiversity-area relationships can be extended from species counts<sup>2,3,27</sup> to higher 233 234 levels of network complexity. The increase in the number of interactions in which each species 235 is involved when area increases indicates that trophic interactions might be more vulnerable to habitat loss than species richness<sup>41,42</sup>. Therefore, the consequences of anthropogenic habitat 236 237 destruction may extend from species loss to wider simplification of natural communities<sup>43–45</sup>, with further consequences for the functioning of ecosystems<sup>14–16,46,47</sup>. The systematic scaling of 238 239 network complexity with area suggests that trophic interactions will be lost with habitat 240 destruction in a predictable manner, and describing the factors shaping this sequence allows us 241 to better anticipate the effects for ecosystem functioning. Importantly, the scaling of the number 242 of interactions with species richness proved variable across datasets, and our null model 243 analysis indicated that other complexity metrics cannot be predicted from species richness

alone. Both of these patterns suggest that we need network-specific knowledge on the
relationship between links and species to accurately predict the effects of habitat loss on
network complexity. To estimate the spatial scaling of the number of links, a general
understanding of the species-area relationship may still suffice, if it is supplemented by
informed assumptions regarding the scaling exponent of the links-species relationship. For such
assumptions, our current results will form a useful point of departure, given their foundation in a
wide set of scales and biomes.

251 Future studies should put effort towards unravelling the potential ecological 252 mechanisms affecting the slopes of NARs, as well as disentangling the importance of the 253 methodological aspects (Supplementary Figure 7 and Supplementary Figure 8; Supplementary 254 Table 7), as has been widely done for SARs<sup>27</sup>. Explicitly investigating the effect of habitat 255 heterogeneity, for instance, would help elucidate the effect of area, not only on network size and 256 the number of links, but also on its modular structure. Such insight is urgently needed, given 257 that heterogeneous landscapes are likely to promote the emergence of compartments within networks<sup>48,49</sup>. Similarly, it is important to investigate how our observations could inform 258 259 predictions about the effects of habitat loss caused, for instance, by link extinction debt or 260 changes in trophic regulation<sup>41,42</sup>. Yet, it is also fundamental to understand how our predictions 261 of the effects of habitat loss on network structure would resemble those obtained by direct 262 experiments of habitat destruction. Gaining a deeper understanding of the mechanisms behind 263 NARs will allow us to not only anticipate the potential consequences of habitat loss, but to also 264 provide management recommendations with a solid foundation in the structure and functioning 265 of natural ecosystems.

# 266 METHODS

267 To measure network-area relationships (NARs) for ecological communities from different

268 biomes across the world, we used 32 empirical datasets comprising species recorded in different

269 localities, and with different types of interactions. This is the most comprehensive synthesis of

270 spatial interaction network datasets to date. Depending on the study, local observations span

271 either a regional or biogeographical spatial extent. Our analysis allowed us to identify 272 universalities in the ways that network properties change with area, for datasets comprising 273 different interaction types. To evaluate whether there are universalities in the ways that 274 networks of ecological interactions scale across space, we quantified the exponents of the 275 relationships between all network properties analysed with area when fitted to a power function. 276 The spatial scaling of network degree distributions was assessed by fitting four different 277 distributions at each spatial domain and comparing the best fit across the entire range of areas. 278 Additionally, we used two different null models to disentangle the effects of the spatial scaling 279 of species richness on the patterns observed in other complexity metrics.

280

# 281 Data classification

32 datasets comprising species distributions and their interactions were obtained from different sources and classified into two categories according to the geographical extent covered: regional or biogeographical spatial domains. The main features of the type of data considered in each category, including sampling method, location, and interaction types, are briefly described below. Additional details about each dataset can be found in Supplementary Text 1.

287

288 Regional spatial domain. Datasets in this category are distinguished from biogeographical 289 networks by their geographical extent and sampling methods. Regional-domain data were 290 collected with locally replicated samples over relatively narrow spatial extents up to ca. 1000 291  $km^2$ . Sample units are generally the same size or a series of localities of roughly equal size. A 292 total of 19 datasets were considered in the regional domain, covering different ecosystems and 293 geographical locations across the globe. The datasets are: networks describing plant-pollinator 294 interactions and host-parasitoid interactions from Mediterranean scrublands located in Garraf (3 datasets<sup>50,51</sup>), and forests located in the Natural Parks of Montsenv (1 dataset<sup>52</sup>) and Olot (1 295 296 dataset<sup>53</sup>) in Catalonia, Spain; host-parasite interactions networks describing insect herbivores 297 (gallers and leaf miners) of the Pedunculate Oak (Quercus robur) and their parasites in a temperate forest in Finland (1 dataset<sup>54</sup>); plant-pollinator interactions from a temperate forest in 298

Argentinian Patagonia (1 dataset<sup>55</sup>); soil food webs from agroecosystems across The
Netherlands (7 datasets<sup>56</sup>); terrestrial food webs within small islands of a temperate saltmarsh
mudflat in south-eastern England (1 dataset<sup>57</sup>); intertidal food webs from the Northeastern
Pacific in Alaska, USA (1 dataset<sup>58</sup>); networks of plant-pollinator interactions and hostparasitoid interactions in fragmented calcareous grasslands in central Germany (2 datasets<sup>59</sup>);
and networks of plant-leafminer-parasitoid interactions from forest fragments embedded in an
agricultural matrix landscape in central Argentina (1 dataset<sup>60</sup>).

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307 Biogeographical spatial domain. Datasets from this category span broad (i.e. biogeographical) 308 spatial extents. There are two types of biogeographical datasets: can be comprised of either (1) 309 datasets where local communities are built based on the information of the presence of all 310 species found in each location and where the interactions between species have been inferred 311 from literature review and expert knowledge; and (2) datasets where local communities are built 312 based on direct observations of ecological interactions at each locality. The datasets of type (1) 313 comprise: the European terrestrial vertebrate trophic network (food web, <sup>61</sup>) divided among 10 314 biogeographical regions that characterise each region according to its climatic conditions. Local 315 communities for this dataset are defined at the  $10 \text{km} \times 10 \text{km}$  grid level (10 datasets); and the terrestrial vertebrate food web of the Catalan Pyrenees (1 dataset;<sup>12</sup>), where local communities 316 317 are also defined at the  $10 \text{km} \times 10 \text{km}$  grid level. The datasets of type (2) comprise networks of 318 plant-herbivore interactions and herbivore-parasitoid interaction sampled from plants of the 319 genus Salix and spanning a large latitudinal gradient from Italy to northern Norway (2 320 datasets;<sup>62</sup>).

321

### 322 Building network-area relationships

To analyse the spatial scaling of network structure we built NARs for each dataset described
above. The procedure to generate NARs was slightly different between spatial domains.

326 Regional spatial domain. The spatial extent was rescaled for each dataset. The smallest spatial 327 scale considered was a single sampling unit. The spatial scaling of network structure was 328 defined by sequentially aggregating each of the sampling units available until all replicated 329 samples were considered. The aggregated sampling units are not contiguous, given that the 330 specific location of each of them was not considered. We analysed network structure at each 331 step of the aggregation procedure. Given that the order in which sampling units (i.e. local 332 communities) are aggregated might generate a bias in the aggregation procedure, we replicated 333 the procedure 100 times for each dataset where the aggregation order was randomly generated 334 without replacement.

335

336 Biogeographical spatial domain. The fundamental difference between the two spatial domains 337 considered is that biogeographical data span large spatial extents, thus comprising communities 338 exposed to large environmental gradients. This environmental heterogeneity suggests potential 339 differences in scaling patterns, when compared to communities sampled regionally, which cover 340 smaller geographical extents. Due to the large spatial extent covered by these datasets, an 341 aggregation procedure where each aggregated sampling unit is randomly selected in space, 342 independently of its location, would generate a fast accumulation of species and links at small 343 areas due to the large heterogeneity among the sampling units (Supplementary Figure 8 and 344 Supplementary Table 7), generating a fast accumulation of species and links at small areas. For 345 this reason, we employ an aggregation method based on neighbouring cells, while still 346 preserving the statistical power of randomising across replicated aggregation instances (here, we 347 refer the reader to Supplementary Figure 8 and Supplementary Table 7 for results based on the 348 random aggregation of sites). Two types of data comprise this category: those for which species 349 presences and their interactions were recorded in the field from single georeferenced locations across the latitudinal gradient<sup>62</sup>; and those where species presences were extracted from species 350 351 distributions maps and interactions were inferred based on the joint condition of species co-352 occurrence in space and an indication that the two species interact, as extracted from the literature<sup>61</sup>. The latter datasets, therefore, include all potential interactions between species, 353

354 which are not necessarily realised locally. Local sampling units of these latter datasets were 355 defined as the 10×10 km cells on a gridded map. The spatial scaling of network structure is thus 356 simulated by merging adjacent cells of the map in an increasing manner. Starting from a 357 randomly chosen cell, subsequent communities occupying larger areas were defined as 358 collections of neighbouring cells forming increasingly larger clusters of cells, chosen in counter-359 clockwise fashion from the starting cell (i.e., spiral fashion). Thus, the sampling units that are 360 aggregated are contiguous. We measured network properties at each step of the aggregation 361 procedure. We repeated the aggregation procedure 100 times. For datasets with georeferenced 362 locations, the spatial scaling of network structure was defined by sequentially aggregating the 363 closest communities (in terms of Euclidean distance) to the starting point.

364

# 365 Network properties

We analysed the spatial scaling of network complexity at both spatial scales measuring the number of species (S), the total number of links (i.e. biotic interactions; L) present in the network and the number of links per species (L/S). We also quantified the relationship between species richness and the total number of links in the network (i.e. links-species scaling).

370 Specialisation was measured at the network level by computing the mean in-degree (mean

 $\label{eq:2.1} 371 \qquad \text{number of resources utilized by a consumer; $L/S_c$) of the community (also known as generality $L/S_c$) of the community (also known as generality $L/S_c$) and $L/S_c$ and $L/S_c$ are supported with the set of the community (also known as generality $L/S_c$) and $L/S_c$ are supported with the set of the set of the community (also known as generality $L/S_c$) and $L/S_c$ are supported with the set of the set of$ 

372 or diet breadth in food webs).

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To further assess how the structure of ecological networks scales across space in different

375 ecosystems, we examined their degree distribution. The degree distribution P(k) of a network is

defined as the probability of finding a species that has *k* links to resources in the network.

377 Degree distributions provide a notion of how links are structured across the network, and their

378 shape has been related to different aspects of community stability<sup>46,63–65</sup>. In particular, scale-free

379 degree distributions are considered the hallmark of network organisation; conferring stability

380 properties to ecological communities. We fitted four different functions that have been

381 identified as typical of the shapes observed in degree distributions in ecological networks:

power law, truncated power law, lognormal and exponential<sup>63,66</sup>. After fitting these distributions 382 383 to the network of interactions built for each spatial scale analysed, for each dataset, the most 384 parsimonious one as measured by Akaike Information Criterion (AIC) was recorded as the best 385 representation of the data. We did that for each replicate of each dataset, selecting the 386 distribution that was best ranked in most of the replicates. Looking at whether the top-ranked 387 distribution (i.e., the lowest AIC) changes with area provides insights into the scale invariance 388 of this network property. Additionally, we looked at the changes of the parameters involved in 389 each function with area (Supplmentary Figure 4).

- 390
- 391 Network-Area relationships fitting

392 After constructing NARs based on the 100 replicates for each dataset, we analysed the scaling 393 relationships by fitting a suite of functions that have traditionally been used to quantify speciesarea relationships<sup>67,68</sup> (see Supplementary Table 4 for the full set of functions tested). These 394 395 scaling functions incorporate in different ways the network property as the response variable 396 and area as the predictor variable, using a characteristic exponent, or a variation of it, to define the relationship of network properties to area size. For example, the power law  $(N = cA^Z)$ , 397 398 relates N (a network property), to area (A) using a constant c and a scaling exponent (z). 399 Similarity in the fitted parameters obtained for different datasets was then used to evaluate 400 universalities in NARs. Scaling functions in Supplementary Table 4 were fitted to each dataset 401 (after network properties were calculated at each spatial extent) using the sar package in R<sup>68</sup>. 402 We used  $R^2$  and p-values to assess goodness of fit. Comparison across models and selection of 403 the best supported model was done with AIC. Amongst all fitted functions, we focused on the 404 top five models that were best supported and selected (if possible) the best ranked model from 405 the power family to facilitate comparisons among datasets. We note that the functions were 406 fitted to the original data in arithmetic space and that the re-scaling of the properties to start at 0 407 was performed for visualisation purposes alone.

408

409 Null models

410 To understand the contribution of species richness and the number of links for the spatial 411 scaling of the associated network properties analysed we used two null models. For null model 412 1, we checked for each dataset the number of species present in each spatial unit and we 413 randomly picked the same number of species from the corresponding metaweb. We then built 414 the network for those species taking from the metaweb all the interactions present between the 415 selected species. Thus, the resulting assemblages can differ from the original networks in the 416 number of links between species and, in turn, in the associated network properties analysed. In 417 contrast, for null model 2, we checked that the number of species is present in each spatial unit 418 but also the number of links present and randomly distributed those links between the selected 419 species. While networks built with null model 1 can potentially preserve part of the structure 420 from each metaweb given that the selected species preserve the links they have in the metaweb, 421 networks built with null model 2 can be considered random networks given that there is no 422 inherited structure from the metaweb. Thus, null model 1 allows us to determine the 423 contribution of the identity of the species (with their respective links) to the observed patterns, 424 while null model 2, allows us to further test whether there is any contribution of area into 425 network structure beyond species richness and the number of links. For both null models, we 426 generated networks of different sizes by adding the number of species of the subsequent spatial 427 units of each dataset. At each step of species addition, we calculated all network metrics. We 428 replicated the procedure 100 times for each dataset. We evaluate the resulting network-area 429 relationships (where area is the number of spatial units from which we extracted the number of 430 species) by fitting a power function, as we did for the original networks. We finally compared 431 the spatial scaling of network properties with those obtained in the original networks 432 (Supplementary Figure 5 and Supplementary Figure 6; Supplementary Table 5 and 433 Supplementary Table 6). 434 **Code availability.** Custom code used to perform the analyses are available online here: 435 https://doi.org/10.5061/dryad.zcrjdfndg, here: https://github.com/nuriagaliana/Ecologicalnetwork-complexity-scales-with-area or here: https://github.com/mlurgi/global-network-area. 436

437 **Data availability.** All datasets analysed during the current study are available online here: 438 https://doi.org/10.5061/dryad.zcrjdfndg, here: https://github.com/nuriagaliana/Ecological-

439 <u>network-complexity-scales-with-area</u> or here: <u>https://github.com/mlurgi/global-network-area</u>.

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- 455

# 456 Author contributions

- 457 NG, JMM and ML designed research with contributions from all co-authors. NG and ML
- 458 conducted research and analysed the data. JB, LC, BCL, CE, IG, CHC, FJ, DM, CM, SOC, SR,
- 459 AR, ISD, AT, DV, SW, TR and WT contributed the data. JMM, VAB, KC, MJF, SL, KM, AM,
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# 462 **Competing interests**

- 463 The authors declare no competing interests.
- 464 Additional information
- 465 Supplementary information is available for this paper.

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- **Table 1**. Parameter estimates for the fit of the extended power model ( $N = cA^{(zA^{-d})}$ ) to the
- 474 network properties analysed. Mean and standard deviation of *d* and *z* are shown for each
- 475 network property analysed in both spatial domains. The model fit was performed on the original
- 476 data in arithmetic space and not on the rescaled properties.

Network property	Parameter	Spatial Domain	
		Regional	Biogeographical
Species	d	$0.08\pm0.03$	$\textbf{-0.38} \pm 0.78$
	Z	$0.48\pm0.12$	$0.05\pm0.41$
Linha	d	$0.07\pm0.03$	$\textbf{-0.19}\pm0.13$
Links	Z	$0.72\pm0.10$	$0.41\pm0.63$
Linha/Sussian	d	$0.05\pm0.11$	$\textbf{-0.31} \pm 0.57$
Links/Species	Z	$0.26\pm0.10$	$0.08\pm0.11$
Indeenee	d	$0.04\pm0.12$	$\textbf{-0.27}\pm0.22$
indegree	Z	$0.31 \pm 0.13$	$0.07\pm0.19$

493

494	Figures

495 Fig 1. Spatial scaling of network complexity. The relationship of species (a,d), links (b,e), and 496 links/species (c,f) with area for regional (top) versus biogeographical (bottom) networks. For 497 each dataset, each point represents the mean value of the analysed network property across the 498 total amount of replicates in the aggregation procedure, for a given area. For all datasets, all 499 network properties have been re-scaled to start at 0 for visualisation purposes. In the 500 biogeographical domain, three datasets show a linear-concave increase of the number of species, 501 network links and links per species with area, similar to those observed in the regional domain; 502 these differences may be explained by differences in sampling methods among datasets (see 503 Methods and Discussion; Supplementary Text 1). 504 505 Fig 2. Scaling of the number of links with species richness. For networks spanning both 506 regional (a) and biogeographical (b) spatial domains, the number of links scales with species 507 following a power law (Supplementary Table 2). Each point represents the mean values across 508 all replicates in the aggregation procedure. Note that for visualisation purposes, all x- and y-509 values in all datasets have been re-scaled to start at 0. To allow convenient comparison with the 510 constant connectance hypothesis and the link-species scaling law, the slope of each log-log 511 relationship is provided in Supplementary Table 2. Note that for constant connectance, the slope 512 equals 2, as the number of links in a web increases approximately as the square of the number of 513 trophic species:  $L \approx S^2$ . For the link–species scaling law, the slope equals 1, as the number of 514 links per species in a web is constant and scale invariant at roughly two:  $L \approx 2S$ ). 515

Fig 3. Spatial scaling of mean indegree and network degree distribution. The mean number of resources per consumer (i.e., mean indegree) increases with area within both regional (left) and biogeographical (right) spatial domains (a-b). Each point represents the mean value across every replicate in the aggregation procedure at a given area. Note that for visualisation purposes, all *x*- and *y*-values in all datasets have been re-scaled to start at 0. The shape of the network

521 degree distributions is consistent across area at both regional and biogeographical domains. In 522 (c-d) two datasets are shown as illustration: Garraf-PP and Galpar, respectively (see 523 Supplementary Text 1). The cumulative probabilities of finding a species in the network that 524 has k links to resources in the network, are normalized by the mean number of links per species 525 in the network. One replicate for a subset of areas is shown for each dataset to facilitate 526 visualization. Each colour represents network degree distribution for a given area. From dark 527 blue, representing the smallest area (i.e. 1 spatial unit), to yellow, representing the largest area 528 for each dataset (40 spatial units for Garraf-PP and 373 spatial units for Galpar). Notice that 529 the starting point of each distribution changes with area, indicating that at smaller spatial scales, 530 the most specialized species of the network have more interacting partners than at larger spatial 531 scales. Yet, the shape of the degree distribution is preserved. The top-ranked model describing 532 the degree distribution of each ecological network across the area range (e-f). Although the 533 degree distribution of most ecological networks was characterised by the same function along 534 the range of areas, the specific shape of each function changed with area (see Supplementary 535 Figure 4). Area values were re-scaled between 0 and 1. 536

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