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# Variational Data Assimilation to Improve Subsurface Drainage Model Parameters

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#### 9 Abstract

Variational data assimilation (VDA) has been implemented to enhance the estimation of the 10 11 unknown input parameters of a new agricultural subsurface drainage model (SIDRA-RU) through assimilating drainage discharge observations. The adjoint model of SIDRA-RU has 12 been successfully generated through the generic automatic differentiation tool (TAPENADE). 13 First, the adjoint model is used to explore the local and global adjoint sensitivities of the 14 15 valuable function defined over the drainage discharge simulations with respect to model input 16 parameters. Next, the most influential parameters are estimated by applying the Variational DA approach. The performed sensitivity analysis shows that the most influential parameters 17 on drainage discharge are those controlling the dynamics of the water table; the second most 18 19 influential parameters manage the drainflow start of each drainage season. Compared to an alternative gradient-free calibration performance, the estimation of these governing 20 parameters by the variational method improves the overall quality of the drainage discharge 21 prediction, in particular in terms of the cumulative water volume. Improved parameters 22

generate less than 5 mm (1%) of the discrepancy between simulated and observed water 23 volumes, based on the five years of daily discharge observations on the Chantemerle 24 agricultural parcels (36 ha). Preliminary numerical tests have shown the potential presence of 25 multiple local minima, thus pointing out the equifinality issues. The latter can be highlighted 26 by the self-compensation of both the physical soil parameters and the main conceptual 27 parameters. For improving the robustness of the parameter estimates, a novel hybrid 28 29 "Bayesian Variational" method is suggested. This method is based on the Bayesian averaging of an ensemble of optimal estimates. 30

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Keywords: agricultural drainage model, adjoint sensitivity analysis, variational data
 assimilation, Bayesian estimation.

#### 34 1. INTRODUCTION

Agricultural drainage systems are designed to remove the excess water within soils subjected 35 to frequent and continuous waterlogging, in particular during the winter season. The 36 agricultural subsurface drained areas contribute to 10% of French arable soils (Vincent, 37 2020). The main objectives of agricultural drainage are to ensure better crop yields 38 39 (Mohtadullah, 1990) and facilitate grazing and access to cultivated fields by farmers. Subsurface drainage flow only occurs when the perched water table rises above the drain 40 depth, as identified by "drainage season". Three drainage patterns can be distinguished during 41 an agricultural drainage season (Lesaffre, 1988): (1) the "pre-flow drainage stage", consists 42 of a gradual saturation of the soil and a limited flow into drains; (2) the "intensive drainage 43 stage" (IDS) featuring a continuous and rapid reaction to each rainfall event; (3) the "drying 44 45 stage". The latter is characterized by the water flow decrease until the drainage flow stops during summer. 46

The hydrological behavior of agricultural drainage networks is highly dependent on 47 hydrometeorological data and soil hydro-physical properties. To understand the hydrology of 48 drained agricultural lands, several numerical models have been developed to approximate the 49 underlying dynamics and physical processes. The appropriate modeling of the perched water 50 table formed in drained soils is complex, time-consuming and still require multiple input data 51 and parameters. In contrast, some simplified models simulate the outputs from a limited input 52 dataset but are not able to describe the physics of hydrological and hydraulic processes 53 (Tournebize et al., 2004; Henine et al., 2010; Henine et al., 2014). 54

Several models are reported in the literature (Gurovich and Oyarce, 2015) for designing 55 drainage networks and monitoring hydraulic patterns. DRAINMOD (Skaggs et al., 2012), 56 EnDrain (Valipour, 2012), CSUID (Alzraiee et al., 2013), and SIDRA that was initially 57 developed at INRAE (formerly Cemagref) by Lesaffre (1988) describe the hydrological and 58 59 hydraulic behavior of drained agricultural lands. More specifically, SIDRA (SImulation of DRAinage) model simulates water table variations at the mid-drain and calculates discharge 60 at the drainage network outlet. Since that time, several model improvements have been 61 62 proposed by Bouarfa (1995) and Bouarfa and Zimmer (2000), including a better consideration of the evaporation term and the water table shape. Recently, a new conceptual 63 module has been introduced to manage the quantity of water seeping through the vadose soil 64 zone (Jeantet et al., 2021; Henine et al., 2022), referred to as SIDRA-RU (RU "Réserve 65 Utile" for Water Holding Capacity). 66

Furthermore, the model uncertainty management is indispensable to enhance the hydrological prediction (Blöschl et al., 2019). Leaving out the structural uncertainty specific to each model, one can focus on improving the input/parameter estimation. This could be achieved in part by considering the most suitable optimization approach in terms of uniqueness, robustness and stability of the final solution. 72 It is known that long-term discharge predictions are needed to adapt and prepare agricultural lands to cope with the effects of climate change (Jiang et al., 2020; Golmohammadi et al., 73 2021). Unlike the short-range predictions (e.g. days, weeks), which may require the 74 knowledge of the hydrological state at the beginning of the prediction window (e.g. flood 75 forecasting), the long-range predictions are almost fully defined by the driving conditions 76 (anticipated rainfall and evapotranspiration) and the model parameters. That is why this paper 77 78 focuses on solving a classical parameter estimation (calibration) problem for SIDRA-RU model, using observations of drainage discharge recorded before the prediction window. 79

First, the most influential model parameters on model outputs should be identified. This can 80 be achieved through the Sensitivity Analysis (SA), a key preliminary step to the optimization 81 exercise (Stange et al., 2000; Migliaccio and Chaubey, 2005; Razavi et al., 2021). Several 82 approaches can be employed in local and global SA, including variance-based approaches 83 84 (Sobol, 1993; Saltelli, 2002) and derivative-based methods (Cacuci, 1981; Morris, 1991). The adjoint sensitivity analysis involves a simultaneous computation of all sensitivities using 85 adjoint operators within a single model run. In addition, inequality links have been 86 87 established between the derivative-based global sensitivity measures (DGSM) and Sobol's indices, thus yielding SA for a large number of input parameters (Kucherenko, 2009; 88 Lamboni et al., 2013). In the same context, Gejadze et al. (2019) derived a generalized 89 relationship between the global sensitivity indices and the Polynomial Chaos coefficients, 90 making SA suitable for high-dimensional models. 91

Once the most influential parameters are identified, one can estimate their values using the available observations. Several approaches are presently available, ranging from the PAP-GR Michel's calibration algorithm (Michel, 1989) highlighted by Mathevet (2005), to more sophisticated Data Assimilation (DA) methods, such as the Kalman filter and its ensemble

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96 variants, Particle filters, variational Data Assimilation (VDA) and hybrid methods
97 (Abbaszadeh et al., 2019).

In particular, VDA method is preferred in operational geophysical large-scale applications, 98 due to its scalability, robustness and potential to handle nonlinear systems under high 99 uncertainty (Courtier et al., 1998; Fischer et al., 2005; Gauthier et al., 2007). However, 100 references on the application of VDA to hydrological models remain relatively scarce. In 101 recent studies, the optimal estimates of the unknown model variables and/or parameters have 102 103 been obtained via minimization of a cost function using the gradient-based methods (Nguyen et al., 2016; Oubanas et al., 2018; Ghorbanidehno et al., 2020; Jay-Allemand et al., 2020). 104 This procedure requires developing the adjoint counterpart of the model, which could be 105 challenging to obtain in some cases (e.g. presence of non-differentiable operators). 106 Application of filtering methods is natural for solving the state or state-parameter estimation 107 108 problems to unable the short and medium range state forecasting. For the pure calibration 109 problems the 'full sample' data processing approach such as VDA should be preferred, given the adjoint model is available. 110

A common issue when solving ill-posed problems in a variational formulation using the 111 112 gradient-based minimization is the likely presence of multiple local minima in the cost function (Dennis and Moré, 1977). A comparison between different optimization algorithms 113 applied to hydrological models was presented by Arsenault et al. (2014). Some of these 114 methods can be highly sensitive to the choice of the initial guess, known as the "background" 115 in DA or the "prior" in statistics. This finding was supported by Pan and Wu (1998), in which 116 an original approach based on the simplex method was used to reveal and avoid the 117 convergence to local minima. Furthermore, Skahill and Doherty (2006) proposed a simple 118 methodology to address the local minima issue. Based on a local search algorithm, their 119 methodology consists of starting the minimization process with different parameter values to 120

increase the probability of reaching the global minima. The same concept has been deployed 121 when calibrating 3D morphodynamic model data (Shoarinezhad et al., 2020) using the PEST 122 parameter estimation package (Welter et al., 2012), thus showing that the optimization 123 process is substantially affected by the local minima issue. 124

However, the presence of multiple local minima is, essentially, a sign of ill-posedness in 125 126 terms of non-uniqueness (equifinality thesis, Beven (2006)). In these circumstances, one should look for the mean of the posterior distribution rather than for one of its modes. The 127 posterior mean is a natural output of the Bayesian methods. Since the direct use of these 128 methods may be difficult in high dimensions, we suggest a simple hybrid approach referred 129 below as the Bayesian Variational (BV) method. That is, the mean is computed over an 130 ensemble of parameter estimates obtained by the variational method, whereas each ensemble 131 member is weighed by the corresponding likelihood. The method is described in detail in 132 133 Section 2.4.2.

134

#### 2. MATERIALS AND METHODS

This paper focuses on improving the long-range predictive performance of the SIDRA-RU 135 136 model (Henine et al., 2022) that runs with a daily time step and requires knowledge of the rainfall, potential evapotranspiration and a set of model parameters. First, the adjoint local 137 and global sensitivity analysis are performed, in accordance with the methodologies 138 139 presented in Goutal et al. (2018) and Gejadze et al. (2019). Next, the VDA method is applied to improve the estimation of the most influential model parameters. Therefore, the uniqueness 140 (equifinality) of the parameter estimation problem is investigated by running the ensemble of 141 142 minimizations from different starting points. Then, the hybrid BV method is performed to improve the robustness of the estimates. A flowchart is embedded below to indicate the 143 connections between the methodology elements (Figure 1). 144





#### 147 **2.1. Study area description and data**

The Chantemerle agricultural sub-catchment (36 ha) has been considered for this study. The field includes a set of artificially drained agricultural parcels. On average, the drains are laid at a depth of 0.9 m and spaced by 10 to 12 meters. The study area is located in Aulnoy (*"Seine et Marne"* department; 70 km east of Paris), within the Orgeval catchment (Tallec (2012); Figure 2).



Figure 2. The "Chantemerle" agricultural field and its general location (48°50' N, 3° 6' E),
"Seine et Marne department", France)

The study site experiments are operated by the GIS-ORACLE research observatory 156 (https://gisoracle.inrae.fr/). The soil is luvisol type (FAO, 2006; Tournebize et al., 2015) with 157 a silty texture (mixture of sand and clay), which is belong to the Aqualf suborder following 158 the US taxonomy system. This type of soil gives rise to a temporary water table, causing fast 159 and continuous water flows during wet periods. The outlet of the subsurface drainage 160 network has been monitored by the water level measurement device (SE-200 OTT). Then, the 161 water level has been transformed into discharge value using calibrated rating curve. The 162 rainfall has been measured near the study site using a rain gauge. Daily potential 163 evapotranspiration (PET) data have been extracted from the SAFRAN database (Vidal et al., 164 2010). PET values have been calculated according to the Penman-Monteith formula. 165 Available data cover September 1<sup>st</sup>, 2008 to August 31<sup>st</sup>, 2013 (Figure 3). 166

Among all recorded data, two wet hydrological years, 2010/2011 and 2012/2013 can be distinguished, accumulating respectively 172 mm and 217 mm of drained water. The three remaining dry years account for a combined 209 mm, nearly equal to the drained water

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accumulated in just a single wet season. This contrast in the quantity of drained water (i.e.



succession of dry and wet seasons) lends more reliability to our results.

Figure 3. Five years (2008-2013) of daily drainage discharge data (Q, black) and rainfall (P,
blue) recorded at the "Chantemerle" field scale. The related daily potential evapotranspiration
values (PET, green) are extracted from the SAFRAN database.

#### 176 2.2. The SIDRA-RU model

SIDRA-RU is a semi-analytical and semi-conceptual model, which is run at the scale of
agricultural plots with a daily time step. The model is based on three coupled modules
(Figure 4; Henine et al. (2022)):

The first module estimates the net infiltration based on more realistic
 evapotranspiration, referred to herein as corrected evapotranspiration (CET). This
 variable takes into account the potential evapotranspiration rate (PET), the available
 soil water content in the soil reservoir, and the crop coefficient (β). The net infiltration

is calculated as the difference between the precipitation (P) and the CET (AppendixA; formulas A.1).

Based on a conceptual reservoir approach, the second module calculates the net 186 recharge to the perched water table as a function of the actual water level in the soil 187 reservoir (S(t)). Three distinct soil reservoir thresholds are accordingly defined 188 namely minimum (S<sub>min</sub>), intermediate (S<sub>inter</sub>) and maximum (S<sub>max</sub>). When S(t) lies 189 between S<sub>inter</sub> and S<sub>max</sub>, the model splits the net infiltration rate into two components 190 via partition parameter ( $\alpha$ ). The first part ( $\alpha$ ) recharges the water table (R) while the 191 second part  $(1-\alpha)$  continues to fill the soil reservoir. When S(t) reaches the maximum 192 level (S<sub>max</sub>), the entire net infiltration recharges the water table (Appendix A; formulas 193 A.2). It is worthwhile to note that S(t) level at the beginning of the hydrological year 194 (N) is based on the soil saturation degree of the year (N-1). 195

The third module simulates drainage discharge (Q) and the mid-drain water table
height (h) above drain by applying the SIDRA model (Lesaffre, 1988). Note that Q
and h values are calculated considering the Hooghoudt equation (Hooghoudt, 1940)
(Appendix A; formulas A.3).



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201

**Figure 4.** Diagram presenting the three SIDRA-RU modules (Henine et al., 2022)

SIDRA-RU inputs are precipitation (P) and potential evapotranspiration data (PET), both considered as forcing variables, along with the model parameters to simulate the mid-drain water level variations (h) and drainage discharges (Q) outputs (Figure 4). The model contains two state variables h(t) and S(t), and eleven parameters.

The parameter set of the model is listed in Table 1. This set includes the SIDRA model 206 207 parameters ( $K_{sat}$  and  $\mu$ ) and the two soil reservoir parameters, i.e., intermediate soil reservoir level  $(S_{inter})$  and reservoir storage capacity during the intensive drainage season (SSDI) 208 which is equal to the difference between  $S_{max}$  and  $S_{inter}$ . The remaining parameters are 209 considered as either given invariants, regardless of field characteristics, or typical invariant 210 characteristics at the field scale. It should be pointed out that the presence of an impermeable 211 layer, on which the drains are laid, limits deep infiltrations (aspect not considered herein). In 212 addition, Augeard et al. (2005) showed that surface runoff becomes negligible under the same 213

214 field conditions since the agricultural plots are being extensively drained, which leads to infrequent soil overflow.

216

215

#### Table 1. Classification of SIDRA-RU parameters

Туре	Parameters	Unit	Notes
SIDRA model	Saturated hydraulic conductivity (K <sub>sat</sub> )	m/day	Measurable but often
	Drainage porosity (µ)	(-)	inaccessible
	Water table shape coefficients (A and C)	(-)	Fixed model parameters
Drainage design	Parcels area (S)	m²	Values usually known
	Drain spacing (l)	m	and dependent on the studied field
	Drain depth (P)	m	
Evapotranspiration	Crop coefficient (β)	(-)	Internal parameters
correction			
Soil reservoir	Water share coefficient ( $\alpha$ )	(-)	
	Intermediate threshold (S <sub>inter</sub> )	mm	Not measurable at the
	Reservoir storage capacity (SSDI)	mm	field scale

#### 2.3. Adjoint sensitivity analysis 217

In addition to the climate data of rainfall and potential evapotranspiration, SIDRA-RU 218 219 features 11 input parameters (Table 1), some of which are difficult to assess directly, such as SIDRA and soil reservoir parameters. In order to reveal the most important parameters to be 220 221 estimated, a sensitivity analysis (SA) is initially performed. Next, the chosen parameters form the control vector to be estimated using the data assimilation method (see Section 2.4). 222

The present paper investigates both local and global adjoint sensitivities. The latter are computed using the derivative-based approaches involving the SIDRA-RU adjoint model to evaluate the gradient of a valuable function J, defined over the model response. Note that the formulation of J depends on the purpose of the study.

The adjoint model is generated using the automatic differentiation (AD) tool TAPENADE 227 (Hascoet and Pascual, 2013), developed at Inria (https://team.inria.fr/ecuador/en/tapenade/). 228 The AD engine processes the Fortran code which implements the mapping from the model 229 inputs to the valuable function I(U). For the AD procedure to succeed, one has to ensure that 230 all operators of the code are differentiable. This is not always the case, especially with 231 conceptual models. For example, the flow routing scheme originally used in AIGA model 232 (described in Jay-Allemand et al. (2020)) was not differentiable. Since the need for a more 233 234 accurate routing scheme was obvious, the new one has been developed taking into account the differentiability requirement. Concerning SIDRA-RU model, a brief inspection of the 235 model equations (see Appendix A) shows that it is conditionally differentiable. Here, 236 'conditionally' means that all parameters are explicitly involved inside the continuous 237 operators, so the gradient with respect to all parameters does exist. However, some 238 239 parameters and state variable S(t) are involved as arguments in the logical expressions (ifthen/elseif-then/endif), which are not differentiable. This means that the computed gradient is 240 241 approximate since it does not include the component associated to the unaccounted graphs. In 242 practice, this component is usually quite small in comparison to the explicit component and, therefore, does not affect the minimization result significantly. To check this, any logical 243 expression can be substituted with the numerical one using the logistic function H(x) =244  $1/(1 + e^{-2kx})$ , where k is a big enough real. Since the exponent is a computationally 245 expensive function, one has to keep a compromise between the gradient accuracy and the 246

computational cost in the modified version of the code. Concerning SIDRA-RU model, no
 modifications have been considered as necessary.

#### 249 2.3.1. Local adjoint sensitivity analysis (LASA)

250 Let us consider a model *M* that transforms input vector *I* into model output vector *X*:

$$M(I) = X \tag{1}$$

In the particular case of the SIDRA-RU model, input vector I includes the vector of parameters  $U_i$  and the climate forcing datasets (i.e. rainfall and potential evapotranspiration):

254 
$$I = \{P(t), PET(t), U\}$$
 (2)

In this study, rainfall P and evapotranspiration PET are considered to be known with high accuracy; thus, their values are not estimated. The input parameters vector U is defined as follows:

$$U = \{S, l, K_{sat}, \mu, A, C, P, S_{inter}, SSDI, \alpha, \beta\}$$
(3)

259 Where: 
$$SSDI = S_{max} - S_{inter}$$

260 The model response *X* represents the simulated temporal evolution of drainage discharge *Q*:

261  $X = Q(t) \quad t \in [0,T]$ (4)

Since in the current SA framework the model output is directly and completely observed, no observation operator is required, and the corresponding observations are hereafter denoted  $X^*$ .

In the local SA context, the sensitivity of the model response with respect to inputs is quantified locally around a given point. Since the true value of U is unknown, it is typically substituted by a prior  $U_b$ , where index "b" stands for the "background" information in DA. Reasonable initial guess of *U* can be provided by an expert evaluation or through introducing
a pre-calibration step.

Adjoint SA method computes the sensitivity of the model response X with respect to the model inputs U via the gradient of the valuable function J(X(U)) := J(U) at the point  $U_b$ (background value):

273 
$$\frac{\partial J(U_b)}{\partial U} = \left(\frac{\partial J(U_b)}{\partial U_1}, \dots, \frac{\partial J(U_b)}{\partial U_n}\right)$$
(5)

All components of the gradient are computed simultaneously using a single adjoint model run, which is the main advantage of the adjoint approach as compared to the finite-difference or statistical methods. Moreover, the derivative obtained is exact (up to the machine accuracy, if all operators involved are differentiable), rather than a finite-difference approximation.

Next, the gradient components must be properly scaled to construct the local sensitivities meaningful in terms of the Analysis of Variance (ANOVA) approach. Let us consider a Gaussian perturbation  $\xi \sim N(0, B)$ , where  $B = diag(\sigma_i^2)$ , i = 1, ..., N and  $\sigma_i$  is the standard deviation. It can be seen that

283 
$$Var[J] = E\left[\left(J(U_b + \xi) - J(U_b)\right)^2\right] = \sum_{i=1}^N \left(\sigma_i \frac{\partial J(U_b)}{\partial U_i}\right)^2 \tag{6}$$

It follows from the above formula that each component of the gradient  $\partial J(U_b)/\partial(U_i)$  has to be scaled by the corresponding  $\sigma_i$ .

In this study, the Kling-Gupta Efficiency (KGE, Gupta et al. (2009)) has been chosen as a valuable function J(U). KGE is widely used to assess the performance of hydrological models and improve simulation quality (Pechlivanidis et al., 2011; Patil and Stieglitz, 2015; Haas et al., 2016; Santos, 2018). The KGE metric ranges from  $-\infty$  to 1, where 1 corresponds to the best match between simulations and observations.

291 
$$J(X, X^*) = 1 - \sqrt{(r-1)^2 + (d-1)^2 + (m-1)^2}$$
(7)

where:

293 - 
$$m = \frac{\overline{X}}{\overline{X^*}}$$
: the ratio between the mean of the simulated and observed data.

294 
$$-r = \frac{\sum_{i=1}^{n} (X_i - \overline{X}) (X_i^* - \overline{X^*})}{\sqrt{\sum_{i=1}^{n} (X_i - \overline{X})^2} \sqrt{\sum_{i=1}^{n} (X_i^* - \overline{X^*})^2}}$$
: the Pearson correlation coefficient.

295  $-d = \frac{\overline{X^*} \sqrt{\sum_{i=1}^n (X_i - \overline{X})^2}}{\overline{X} \sqrt{\sum_{i=1}^n (X_i^* - \overline{X^*})^2}}$ : the ratio between the standard deviation of simulated and

observed values.

Moreover, the use of KGE allows the performance of adjoint-based and Sobol methods (Sobol, 1993) to be compared. The latter had recently been explored in (Henine et al., 2022).

#### 299 2.3.2. Global adjoint sensitivity analysis (GASA)

While the local SA answers how a small perturbation around the background value  $U_b$ 300 301 influences the model response X, the global SA focuses on the model output variance or, 302 more precisely, on how the input variability influences the output variance. The global SA reveals which parts of the output variance are due to different inputs by estimating Sobol 303 indices, which are a central tool (and key point) since they provide a quantitative and rigorous 304 overview of the influence of inputs on the model response. The commonly used global SA 305 method is the variance-based ANOVA decomposition. This paper introduces the adjoint-306 based global SA described in (Gejadze et al., 2019), featuring a methodology that derives a 307 generalized relationship between the global sensitivity indices and the Polynomial Chaos 308 coefficients. 309

The variability of inputs is represented by a sample of random vectors from a given probability distribution. For example, one can consider the Gaussian distribution using the background vector  $U_b$  as its mode. Since in our case the informative priors for U are either not available or not accurate enough, we use the uniform distribution to define a random sample within a defined interval  $[a_i, b_i]$  for each element of U, as follows:

315 
$$\begin{aligned} \xi_i \sim u(a_i, b_i) \\ \xi_i = (b_i - a_i) * \varepsilon_i + a_i \end{aligned} , \quad \varepsilon_i \sim u(0, 1) \end{aligned} \tag{8}$$

The bounds *a* and *b* are specified for all SIDRA-RU parameters. The choice of the uniform distribution overcomes the need for background  $U_b$  and, subsequently, for the pre-calibration step.

The sample size is chosen by taking into account the computational requirements and model complexity into account. Since the SIDRA-RU is relatively inexpensive, it is feasible to use a relatively large sample size, which can be chosen to compromise between the stability of the sensitivities and CPU time.

In this paper, we only consider the first-order 'main-effect'  $S_i$  indices as well as the 'totaleffect'  $S_i^T$  indices (Saltelli, 2002), defined as follows:

325
$$\begin{cases}
S_{i} = \frac{Var(E[J(\xi)|\xi_{i}])}{Var(J(\xi))} \\
S_{i}^{T} = \frac{Var(J(\xi)) - Var(E[J(\xi)|\xi_{-i}])}{Var(J(\xi))}
\end{cases}$$
(9)

Where  $J(\xi)|\xi_i$  represents the random value of *J*, with the input component  $\xi_i$  being fixed at its generic value;  $\xi_{-i}$  stands for 'all, but i', and  $Var(J(\xi))$  is the total variance of  $J(\xi)$ .

While the main-effect indices, which are commonly used in ANOVA analysis, quantify the single influence of individual variables or some groups of variables, the total-effect indices measure the influence of a variable jointly with all its interactions with other variables. Alternatively, these indices can be viewed as a measure of the variability remaining to the output when all but one input variables are fixed. If the total sensitivity index of a variable is sufficiently small, this variable can be removed from further analysis, because neither the variable nor any of its order interactions have an impact on the valuable function. Thus, the total sensitivity index can be used to detect the essential variables for the calibration process.

The methodology presented in Gejadze et al. (2019) suggests a generalized approximation for the upper bound of the total-effect indices  $S_i^T$  based on a computation of the derivative of the valuable function *J*. We limit ourselves here to the first-order analysis. Future work will investigate the approximation of  $S_i^T$  in considering second-order interactions. Given that  $\xi$  is a vector constituted of independent and identically distributed normal random variables, the following estimate holds:

342 
$$S_i^T Var(J(\xi)) \le E\left[\left(\frac{\partial J(\xi)}{\partial \xi_i}\right)^2\right]$$
(10)

In the case of independent uniformly distributed variables defined over the interval  $[a_i, b_i]$ (Lamboni et al., 2013), the following can be written:

345 
$$S_i^T Var(J(\xi)) \le E\left[\left(\frac{\partial J(\xi)}{\partial \xi_{in}}\right)^2\right] \left(\frac{(b_i - a_i)}{\pi}\right)^2 \tag{11}$$

Similarly, we employ the same formulation of the valuable function introduced in Section 2.3.1. The gradient of J within the GASA framework is computed using automatic differentiation. A detailed algorithm of the GASA applied to the SIDRA-RU model is described in Figure 5.



350

351

Figure 5. Simplified version of the global adjoint SA algorithm

#### 352 2.4. Variational data assimilation algorithm

Data assimilation methods are widely used in the geophysical sciences (Reichle, 2008). These methods provide the best possible estimates of unknown model inputs (initial and driving conditions, parameters) by combining all available information: observations, *a priori* data (coming from expertise or preprocessing), and knowledge of the physics of the underlying phenomena incorporated into the model.



Figure 6. Scheme describing the operating mode of the direct problem (blue arrows) and
 inverse problem (red arrows) methods for the case of the SIDRA-RU model

358

In the 'forward' or 'simulation' mode, the forcing terms (rainfall and PET) and the SIDRA-RU model parameters are provided as known inputs in order to simulate the drainage discharge. The 'backward' or 'inverse' mode seeks to estimate the unknown input parameters by assimilating available information from the observed output discharge (Figure 6).

The recent paper written by Jeantet et al. (2021) presents an example that uses the "Michel 365 calibration (PAP-GR)" algorithm to estimate the parameters ( $K_{sat}$ ,  $\mu$ , SSDI and  $S_{inter}$ ). This 366 algorithm was originally developed at INRAE (formerly IRSTEA) to calibrate the GR (Génie 367 Rural) hydrological models (Perrin et al., 2003). An operational version of the algorithm 368 (airGR) is available in an R package (Coron et al., 2017). This method is initially based on 369 assigning a probability distribution law for each calibrated parameter in order to target the 370 optimal variation domain. The calibration procedure is then refined by searching for the 371 optimal parameter values within the pre-selected bounds. The objective function used through 372

the PAP-GR routine depends on the calibration target (e.g. low or high flow, volume improvement). For more detail, one can refer to the paper of (Jeantet et al., 2021).

The VDA yields optimal estimates of the unknown parameters by minimizing the cost function

377 
$$J(U) = \frac{1}{2} \left\| \left( R^{-1/2} M(U) - X^* \right) \right\|^2 + \frac{1}{2} \left\| B^{-1/2} (U - U_b) \right\|^2$$
(12)

Under the box constraints  $a \le U \le b$ . Here,  $B = E(\xi_b \xi_b^T)$  and  $R = E(\xi_o \xi_o^T)$  are the covariance matrics of the background and observation errors,  $\xi_b$  and  $\xi_o$ , respectively. As in Section 2.3.1, we consider  $B = diag(\sigma_i^2), i = 1, ..., \bar{n}$ , where  $\sigma_i$  is the standard deviation, and  $\bar{n}$  is the dimension of the control subset. Note that a-priory information on the parameters is represented by the bounds *a* and *b*. Since (12) implies the Gaussian distribution of  $\xi_b$ , we define  $\sigma_i = ((b_i - a_i)/2)/3$ .

The observation error  $\xi_o$  includes directly the measurement error, plus projections into the observation space of orher errors (forcing data errors, structural errors). Since *R* is difficult to evaluate, we assume it is diagonal, with uniform variance  $\alpha^2$ . Then (12) can be written in the form

388 
$$J(U,\alpha) = \frac{1}{2} \| (M(U) - X^*) \|^2 + \frac{\alpha^2}{2} \| B^{-1/2} (U - U_b) \|^2$$
(12a)

Where  $\alpha$  has to be estimated alongside *U* from an auxiliary condition on *U*. This is a 'relaxed' formulation which roughtly takes into account all uncertainties not included into the control vector. In practice, instead of solving this problem explicitly, one can consider the minimization problem for the cost function

393 
$$J(W) = \frac{1}{2} \left\| \left( M(U_b + B^{1/2}W) - X^* \right) \right\|^2$$
(13)

Where *W* is defined via the change of variables  $U = U_b + B^{1/2}W$ , complemented with an appropriate iterations stopping criteria (L-curve). This is an iterative regularization implementation of (12a), as described in (Gejadze and Malaterre, 2017; Oubanas et al., 2018).

In the present study, a limited memory L-BFGS-B algorithm (Zhu et al., 1997) has been 398 employed to perform the minimization step involving box-constrained set of parameters in 399 order to avoid unphysical solutions. As explained above, the choice of the VDA method is 400 largely motivated by the need to investigate and optimize the upgraded version of SIDRA-401 RU (to be reported in a future paper), which will be able to simulate the daily nitrate 402 concentration values by integrating the unknown distributed variable, in addition to the model 403 input parameters. The mentioned distributed variable represents the potentially leachable 404 quantity of the nitrate initially trapped in the soil surface layer. This quantity is estimated 405 once a year, generally at the end of the autumn season. The control vector will then be 406 extended accordingly. For this reason, the choice of variational DA is justified since the 407 algorithm sought needs to be robust to the present uncertainties and heterogeneity of the 408 variables involved. Hence, this paper offers the first implementation of the VDA to SIDRA-409 type models, which will be adapted to the new upgraded version. 410

Let us note that (12) is the strong-constraint variational formulation, which is valid under the 411 perfect model assumption. In the variational DA framework, the model error is usually 412 treated via the weak-constraint formulation (Tremolet, 2006). This approach is proved useful 413 in the classical DA in weather and ocean prediction systems, where the model state at the 414 beginning of the prediction window is of major concern. In hydrology, the method has been 415 used as a part of the hybrid-filtering algorithm in (Abbaszadeh et al., 2019). However, the 416 usefulness of the weak formulation for solving the pure calibration problems is far less 417 evident. Besides, being a conceptual and approximate model, SIDRA-RU is characterized by 418

a structural uncertainty, which is not exactly the model error as it is understood in the
classical applications (for example, due to random forcing). A more appropriate approach to
deal with the structural uncertainty seems the model selection in the ExpectationMaximization framework (Beal, 2003). This is a topic for a possible future development, but
we keep the strong-constraint DA formulation as a first step in this direction.

#### 424 2.4.1. Toward avoiding the local minima issues

Due to the ill-posedness of certain inverse problems, the corresponding minimization process 425 may suffer from the presence of the local minima, which prevent it from reaching the global 426 one. Various approaches have been investigated to mitigate this difficulty. Inspired by the 427 works of Duan et al. (1992) and Skahill and Doherty (2006), we suggest using an ensemble 428 minimization framework. The idea here is to start each individual minimization process from 429 a different first guess, which might result in different minimization trajectories, thus limiting 430 the dependence of the optimal solution from the prior knowledge and its uncertainty. The 431 ensemble of priors/backgrounds (used as initial points for minimization) is generated using a 432 433 uniform distribution and is limited to the sufficient subset elements, as identified through the GASA method (see Section 3.1). Two configurations have been tested herein: 1) the input 434 parameters are perturbed one by one, setting the remaining ones at their optimal values; and 435 2) parameters are perturbed simultaneously. The ensemble size is then chosen to compromise 436 target accuracy and available computational resources. 437

438

#### 2.4.2. Hybrid BV method: description

As mentioned above, if the estimation problem is ill-posed in terms of the uniqueness condition, looking for the global minimum might be useless in principle. That is, the resulting posterior distribution could be multi-modal, with modes having nearly the same probability values, but corresponding to noticeably different combinations of parameters. Since both the 443 model and the observations are not 'perfect', choosing the posterior mode with the largest 444 probability (or the 'global' minimum in terms of the cost-function) could be a wrong 445 decision. Thus, the remedy is to look for the posterior mean instead. This should result into 446 more robust (reliable) parameter estimates. Since the direct use of the Bayesian methods 447 could be computationally too expensive, we suggest a novel hybrid BV method combining 448 variational and Bayesian elements, which can be considered as an upgrade to the method 449 presented in (Skahill and Doherty, 2006).

Let us consider an ensemble of optimal solutions  $U_i$ , i = 1, ..., N, where each  $U_i$  is the result of minimization of the cost-function (13), involving a randomly chosen background  $U_b$  from the probability distribution  $\rho(U)$ . Let  $J_i = J(U_i)$  be the corresponding value of (13), and  $J_{min}$ the minimum value of  $J_i$  over the ensemble. Now we introduce the likelihood function in the form:

455 
$$\mathcal{L}_{i}^{k} = exp\left(-2^{k}\left(\frac{J_{i}}{J_{min}}-1\right)^{2}\right)$$
(14)

456 where k is the parameter, which controls the decay rate of the likelihood function. This 457 parameter is similar to the regularization parameter used in the classical Tikhonov approach. 458 Then, the posterior ensemble mean and variance are computed as follows:

459 
$$\overline{U}^{k} = \frac{1}{c_{1}} \sum_{i=1}^{N} U_{i} \mathcal{L}_{i}^{k} \rho(U_{i})$$
(15)

460 
$$Var(U)^{k} = \frac{1}{c_{1}} \sum_{i=1}^{N} \left( U_{i} - \overline{U}^{k} \right) o\left( U_{i} - \overline{U}^{k} \right) \mathcal{L}_{i}^{k} \rho(U_{i})$$
(16)

461 where 
$$C_1 = \frac{1}{N} \sum_{i=1}^{N} \mathcal{L}_i^k \rho(U_i)$$

462 The L-curve approach is adapted here for choosing k. L-curve is defined as a parametric 463 curve  $\{J(\overline{U}^k), D^k\}, k = -20, ..., 20$ , where

464 
$$D^{k} = (\overline{U}^{k} - \overline{U}^{-\infty})^{T} (Var(U)^{k})^{-1} o(\overline{U}^{k} - \overline{U}^{-\infty})$$
(17)

465 is the probabilistic (Mahalanobis) distance between  $\overline{U}^k$  and

466

$$\overline{U}^{-\infty} = \frac{1}{N} \sum_{n=1}^{N} U_n \tag{18}$$

The idea of the 'L-curve' method is to choose the value of *k* located near the corner, so that both the cost-function  $J(\overline{U}^k)$  and the distance  $D^k$  are simultaneously minimized. One particular example of the L-curve encountered during computations is presented in Figure 13. Note that the uniform prior distribution for  $U_b$  is assumed in this paper, therefore the posterior moments are fully defined by the likelihood function (14).

#### 472 **2.5. Evaluation of the temporal robustness of the model**

An ensemble of operational tests proposed by KlemeŠ (1986) has been widely used to 473 evaluate the spatiotemporal robustness of hydrological models (Henriksen et al., 2003; Guo et 474 al., 2020; Jeantet et al., 2021). In the following, we have selected the split sample test to 475 evaluate the temporal robustness of the SIDRA-RU model. This method divides the study 476 period into two subperiods (e.g.  $P_1$  and  $P_2$ ). The first window  $P_1$  is considered as an 477 assimilation period. Hence, the daily discharge is assimilated using VDA in order to estimate 478 479 the control vector U. The parameter estimates obtained are then validated by comparing the corresponding model predictions with observations over the period  $P_2$ . Next, the assimilation 480 481 and validation periods  $P_1$  and  $P_2$  are swapped to validate the estimator's temporal robustness.

Overall, four metrics have been selected to evaluate the model's predictive performance. One can recall the cost function (Section 2.4; Equation 13) and the KGE metric (Section 2.3.1; Equation 7), relevant to assess the variability of the drainage discharge. In addition, we use the root mean square error (RMSE (Kenney, 1939)), which informs of the residuals spread.

486 
$$RMSE = \sqrt{\frac{\sum_{i=1}^{N} (X_i^* - X_i)^2}{N}}$$
(19)

and the Nash-Sutcliffe Efficiency (NSE (Nash and Sutcliffe, 1970)), as defined in Equation
20, which is introduced to assess the performance of several hydrological and water quality
models.

490 
$$NSE = 1 - \frac{\sum_{i=1}^{N} (X_i^* - X_i)^2}{\sum_{i=1}^{N} (X_i - \overline{X^*})^2}$$
(20)

In addition, we are quantifying the water balance conservation by calculating the discrepancyon cumulative drained water volumes.

#### 493 **3. RESULTS**

#### 494 **3.1. Adjoint local and global sensitivity analyses**

The LASA method was first applied to assess the influence of SIDRA-RU input parameters on the KGE metric. First, the local sensitivity around 'pre-calibrated prior' (Table 2) was computed (as illustrated in Figure 7 with red crosses) using the entire data (Figure 3). Results show that the highest sensitivities are to soil parameters  $\mu$  and  $K_{sat}$  and reservoir parameter  $S_{inter}$ .

Second, in order to investigate the impact of the prior on the resulting sensitivities, a set of 1000 random input values has been generated using the uniform distribution defined over the interval [a,b] (Equation 8 and Table 2). The corresponding sensitivities have been then computed using LASA. The boxplots in Figure 7 show the range of impact of each parameter. The highest range is observed for parameters  $K_{sat}$ ,  $\mu$  and further  $S_{inter}$  A and  $\beta$ whereas the smallest range was identified for parameters S, P and  $\alpha$ .





Further analysis has therefore been performed using the GASA method over the entire range of parameter variation, as defined within bounds [a,b]. A random ensemble of SIDRA-RU input parameters has been generated from the uniform distribution. The ensemble size was chosen to balance the stability of results and low CPU time. Hence, testing different ensemble sizes (ranging from 100 to 10000) allowed the selection of n = 2000, ensuring stable sensitivity results. One can note that only ~2.6 s was needed to complete the 2000 simulations.

Table 2 lists the values of the boundaries (a) and (b) assigned according to both the prior knowledge of SIDRA-RU model parameters and the data collected from the Chantemerle field. Accordingly,  $K_{sat}$ ,  $\mu$ ,  $S_{inter}$ , and *SSDI* parameter bounds have been selected from (Henine et al., 2022). The other parameter boundaries was assigned so as to avoid outliers.

Input	S	I	Ksat	μ	Α	С	Р	SSDI	S <sub>inter</sub>	α	β
	(ha)	(m)	(m/day)	(-)	(-)	(-)	(m)	(mm)	(mm)	(-)	(-)
a	35.40	3.50	0.02	0.010	0.80	0.85	0.60	10.00	55.00	0.00	0.50
b	37.40	6.00	1.00	0.130	0.90	0.95	1.20	55.00	225.00	1.00	1.00
Pre-											
calibrated	36.40	5.00	0.27	0.042	0.86	0.89	0.90	46.48	92.58	0.33	0.90
Prior											

The obtained 'total-effect' sensitivities using the GASA approach are shown in Figure 8. The results indicate that the saturated hydraulic conductivity  $K_{sat}$  is the most influential input parameter, followed by the drainage porosity  $\mu$  and then the intermediate water level threshold in the conceptual reservoir  $S_{inter}$ , followed by the other parameters. This finding means that only  $K_{sat}$ ,  $\mu$  and subsequent  $S_{inter}$  parameters exert the strongest impact on KGE values. Although *SSDI* is classified as a low-impact parameter, it is included in the control vector due to its strong link with  $S_{inter}$  parameter (Section 2.2).





Figure 8. The influence of input parameters on the valuable function J (KGE)

#### 530 **3.2. SIDRA-RU parameters estimation using the VDA method**

Based on the results of GASA (Section 3.1), the VDA method was implemented to estimate the impacting part of the input vector on the model output (i.e.  $U = \{K_{sat}, \mu, S_{inter}, SSDI\}$ ). Two experimental set-ups have been considered:

- a) the split sample test, considering the two-time windows  $P_1$ , covering the period from 2008 to 2010, and  $P_2$  covering the period 2010-2013, to assess the model's predictive performance;
- b) the full sample test (over the entire observation period,  $P_1 \cup P_2$ ), with the resulting simulated discharge to be compared to the one obtained with parameters estimated using the PAP-GR calibration method.

#### 540 3.2.1. Validation of the variational DA framework

Table 3 shows the results of the split sample test. Good performance in terms of KGE, NSE and RMSE is achieved for both the assimilation and validation periods (KGE > 0.66; NSE > 0.53; RMSE < 2.44 l/s). Moreover, swapping the assimilation and validation periods  $P_1$  and  $P_2$  leads to relatively comparable estimates (e.g. S<sub>inter</sub> estimated at 71.2 mm and 85.8 mm over periods  $P_1$  and  $P_2$ , respectively), and similar performance metric values. One can also notice that distances between the SIDRA parameters (e.g.  $\mu$  on P1 vs  $\mu$  on P2) are smaller compared to those for the conceptual parameters (e.g.  $S_{inter}$  on P1 vs  $S_{inter}$  on P2).

Simulations	Estimation period P1 (2008-2010)			Validation period P2 (2010-2013)			Estimated parameters					
	RMSE	NSE	KGE	Cost	RMSE	NSE	KGE	Cost	Ksat	μ	SSDI	Sinter
	( <b>l</b> /s)	(-)	(-)	(-)	( <b>l/s</b> )	(-)	(-)	(-)	( <b>m/d</b> )	(-)	(mm)	(mm)
P1 → P2	1.62	0.56	0.66	0.74	2.44	0.71	0.77	2.50	0.19	0.042	40.8	71.2
	Validation period P1 (2008-2010)			Estimation period P2 (2010-2013)			E	stimated	paramet	ers		
P2 → P1	1.66	0.53	0.70	0.78	2.41	0.72	0.77	2.45	0.23	0.042	35.2	85.8

548 **Table 3.** Split sample test results based on assimilation of the drainage discharge data

#### 549 3.2.2. VDA vs. Gradient-free PAP-GR method

Below, the VDA method is compared to the gradient-free PAP-GR calibration method by estimating the control vector values over the entire observation period (2008-2013). Three experimental configurations are considered with the following sets of input parameters:

- (a) Non-informative (open-loop) inputs (priors/backgrounds)
- (b) Calibrated inputs from Michel's calibration algorithm (PAP-GR)
- (c) Estimated inputs from the VDA method, taking the background values from (a)

Table 4 shows the input parameters obtained for configurations (a), (b) and (c), along with the corresponding performance metrics values. Note that the improvement achieved by VDA compared to the PAP-GR calibration algorithm is indicated in green brackets.

Table 4. Performance criteria resulting from: (a) the use of open-loop parameters; the
parameter estimation using both (b) the PAP-GR and (c) the VDA minimization methods

	С	ontrol ve	ctor value	s	Performance criteria			
	K <sub>sat</sub> µ S		SSDI	Sinter	RMSE	NSE	KGE	
	(m/day)	(-)	( <b>mm</b> )	( <b>mm</b> )	( <b>l/s</b> )	(-)	(-)	
a	0.9	0.06	18.0	30.0	4.66	-0.46	0.08	
b	0.272	0.043	46.48	92.58	2.13	0.70	0.76	
c	0.228	0.044	41.93	84.84	2.07 (-0.06)	0.71 (+0.01)	0.78 (+0.02)	

In the experimental set-up (c), the VDA is used starting from a non-informative background 561 from (a) in order to estimate input parameters  $K_{sat}$ ,  $\mu$ ,  $S_{inter}$  and SSDI simultaneously. Since 562 the bounds of the physical parameters range have not been integrated into SIDRA-RU, a 563 constrained cost minimization is performed using the L-BFGS-B method to eliminate all non-564 physical solutions. Table 5 summarizes the lower and upper bounds assigned to each 565 parameter. The choice of  $K_{sat}$  and  $\mu$  bounds is based on field measurements carried out in 70 566 referenced soils in France (Lagacherie, 1987). SSDI and Sinter bounds are presented in 567 (Henine et al., 2022). The estimated control vector (c) is comparable to that derived from the 568 split sample test in the first experiment (Table 3). 569

570

 Table 5. Assigned lower and upper bound values of each calibrated input parameter

Parameter	K <sub>sat</sub> (m/day)	μ(-)	SSDI (mm)	Sinter (mm)
Lower bound	0.03	0.02	10.0	55.0
Upper bound	1.5	0.07	60.0	225.0

Figure 9 illustrates the observed and simulated discharges, over the total target period, according to the three experimental configurations (a, b, c). In addition, the respective cumulative water volume has been plotted in (d), thus making it possible to compare the performance obtained using the variational calibration method and that resulting from the application of the PAP-GR calibration algorithm.

The model performance after VDA is slightly better in terms of KGE, NSE and RMSE 576 values, as compared to the PAP-GR calibration routine. Excluding the first year of data, 577 generally used as a warming period to adjust the saturation degree in the conceptual reservoir, 578 the discharge values are well represented in both cases (Figure 9b and 9c). In addition, the 579 analysis of the simulated cumulative drained water over the total target period (five drainage 580 seasons) in comparison with the measured discharge, indicates only a 1% ratio between 581 measured and simulated total drained water volume using VDA ( $\Delta V = 4.8 \text{ mm}$  over five 582 years of data; Figure 9d) versus 10% when using the PAP-GR calibration algorithm ( $\Delta V =$ 583 58.7 mm). 584



586 **Figure 9.** Graphical comparison between experimental configurations (a), (b) and (c);

587

(d) shows the corresponding cumulative drainage discharge

#### 588 **3.3. Investigating the local minima issue**

589 The second part of this study focuses on the issue of local minima. Two experimental set-ups 590 have been considered. First, we analyze the convergence process for a given single parameter, assuming that the remaining parameters are known and set to their optimal values.
Second, we investigate the convergence process when all parameters are updated
simultaneously.

594 3.3.1. Single parameter estimation

In this experiment, each parameter  $U_i$  is estimated separately while assuming that all other 595 inputs are known and set to their optimal values (Table 4, Row C). Hence, the control vector 596 includes one parameter at a time. For each parameter  $U_i$ , an ensemble of priors/backgrounds 597 598 of size n=100 is generated and an ensemble estimation using the VDA is performed. Figure 10 illustrates the convergence process for each parameter, starting from different initial 599 values within the given ensemble. The red color represents successful minimization cases, 600 601 where a cost function minimization has been reached. The corresponding ratios are presented in the legend, together with the unsuccessful cases shown in gray. Note that minimum cost 602 603 values are generally reached after only 10 iterations.



604

Figure 10. Convergence trajectories of parameters from prior/background values towards the
 optimal values during the single parameter estimation process

It can be noticed that the ensembles of parameters  $K_{sat}$  and  $\mu$  converge towards unique values of 0.22 m.d<sup>-1</sup> and 0.047, respectively, regardless of the initial background value. However, *SSDI* and  $S_{inter}$  parameters converge towards different optimal values. Those corresponding to the minimum cost function are 42 mm and 82 mm, respectively, which are still consistent with the result obtained in Section 3.2.

#### 612 3.3.2. Simultaneous minimization

Similar to the first set-up, in this second experiment, an ensemble of size n=100 of prior/background vectors is generated. The simultaneous minimization of all parameters was then performed using the VDA method. Figure 11 shows the parameter sets that converge toward minimum cost function values, presented here in red. Four main clusters (red dots) have been identified. Similar cluster positions for the two physically based parameters  $K_{sat}$ [0.19-0.25] m.d<sup>-1</sup> and  $\mu$  [0.040-0.047] can be detected. Interestingly, the conceptual-based parameters *SSDI* [47-61] mm and  $S_{inter}$  [65-81] mm display opposite cluster positions.





621

**Figure 11.** Final cost values according to their corresponding parameter values

Similar to the first experiment, Figure 12 shows the minimization trajectories. The red color refers to the successful convergence cases reaching the minimum cost function value. It can be observed that several solutions are possible when estimating all parameters and no global minimum can be clearly identified. Dashed lines correspond to the optimal solution of each parameter in terms of the minimized cost function. The best control vector estimates corresponds to:  $K_{sat} = 0.22$  m/day,  $\mu = 0.044$  (-), SSDI = 48.68 mm, and  $S_{inter} = 78.76$  mm.



Figure 12. Convergence trajectories from prior/background values towards optimal values
 (case of simultaneous minimization)

#### 631 **3.4. Hybrid BV approach to improve the solution stability**

628

The model parameters are estimated by the hybrid BV method following the "split sampletest" methodology (see Section 2.5).

The parameter estimations presented in Table 3 (for P1 and P2), and Table 4 (for P) have been embedded in Table 6 to allow a better comparison between hybrid DA and variational DA methods. First, one can notice that the difference between the estimates  $K_{sat}$  and  $S_{inter}$ obtained using the hybrid DA on different periods P1 and P2 is less, whereas for  $\mu$  and *SSDI*  is slightly larger. One key model characteristic which defines the water table dynamics is given by  $S_{max} = SSDI + S_{inter}$ . The difference  $S_{max}|_{P2} - S_{max}|_{P1}$  by the hybrid BV method makes 5.29 mm, against 9.1 mm by the variational DA. This indicates that the results by the hybrid BV method, as expected, are generally more robust. Secondly, while the estimates  $K_{sat}|_{P}$  and  $\mu|_{P}$  are nearly the same for both methods, they are slightly different for  $SSDI|_{P}$  and  $S_{inter}|_{P}$ .

644 **Table 6:** Robust SIDRA-RU parameter estimations obtained by applying both, Hybrid BV
645 and VDA methods

Method	Parameter	K <sub>sat</sub> (mm.day <sup>-1</sup> )	μ(-)	SSDI	S <sub>inter</sub> (mm)	S <sub>max</sub> (mm)	k
				(mm)			
Hybrid BV	Period P1 (2008-2010)	0.252	0.049	40.74	71.52	112.26	10
	Period P2 (2010-2013)	0.225	0.043	34.01	83.54	117.55	13
	Full period (2008-2013)	0.227	0.045	50.86	76.67	127.53	15
VDA	Period P1 (2008-2010)	0.190	0.042	40.80	71.20	112.00	-
	Period P2 (2010-2013)	0.230	0.042	35.20	85.80	121.00	-
	Full period (2008-2013)	0.228	0.044	41.93	84.84	126.77	-

As an example, Figure 13 shows the L-curve obtained by processing the ensemble of optimal

solutions corresponding to P2 period. Here, one can see a distinctive corner at k = 13. This

648 position ensures the robustness of the solution by selecting low values of D(k) and J(k).





650 **Figure 13.** L-curve, Chantemerle P2 2010-13, "k" are plotted in dark-red color

#### 651 4. DISCUSSION

The paper's content revolves around improving the new agricultural drainage model. First, the adjoint sensitivity analysis (local and global) has been conducted to identify the most impacting parameters of the SIDRA-RU model. Then, VDA and hybrid BV frameworks have been validated over five years of daily-observed discharge data and used afterwards to estimate the best combination of SIDRA-RU parameters.

#### 657 **4.1. Adjoint SA**

The local adjoint SA result suggests that  $\mu$  and  $K_{sat}$  are the most influential parameters. However, one can notice that the choice of prior could change the parameter ranking which emphasizes the importance of the global analysis. In fact, no priors are required to apply the global adjoint approach. The global SA method indicates that  $K_{sat}$  and  $\mu$  have the most impact on the model output. These two parameters govern the physically based part of the model and describe the dynamics and reactivity of the perched water table and discharges. The conceptual parameter  $S_{inter}$ , ranked third, controls the beginning of the intensive drainage season, which generally occurs during winter, while  $\beta$  is dependent on the type of vegetation cover. Since information on crop rotation is not always available, parameter  $\beta$  is set by default equal to 1. Future work will investigate the possibilities of integrating a crop rotation module should data be available.

The results obtained are consistent with the analysis performed using Sobol's total and main-669 order sensitivity indices in (Henine et al., 2022). This output confirms that  $K_{sat}$  and  $\mu$  have 670 the highest impact on KGE, whereas SSDI and  $S_{inter}$  parameters practically determine the 671 starting date of the drainage season. These four parameters constitute the sufficient control 672 673 subset of parameters, or the 'control vector'. It should be noted that one limitation of the Sobol method is the significant computing resources required (looss and Lemaître, 2015; 674 Zhang et al., 2015). This constraint accounts for the main reason why the analysis was 675 conducted on a limited input vector in (Henine et al., 2022) paper. Such a limitation has been 676 overcome by using the adjoint-based sensitivity analysis approach, which allows considering 677 678 all SIDRA-RU inputs in a single model run. In addition, the limited number of SIDRA-RU input parameters avoids the need to use sophisticated sampling methods, e.g. Latin hypercube 679 680 (Viana, 2016), which are typically reserved for models with higher number of inputs 681 (Manache and Melching, 2004).

682 **4.2. Model parameters improvement** 

The split sample test validates the accurate implementation of the VDA method. Moreover, we compare the estimated values of SIDRA-RU parameters with those found in the literature. Therefore, it has been possible to verify the estimated parameters against more than 100

reported values of  $K_{sat}$  and  $\mu$ , corresponding to the silty soil (consistent with the Chantemerle 686 soil texture). The latter have been extracted by Jeantet et al. (2021) from the drainage 687 reference sector reports (Lagacherie, 1987). The probability density functions (PDF) of the 688 saturated hydraulic conductivity ( $K_{sat}$ ) and drainage porosity ( $\mu$ ) have been calculated using 689 their lognormal distributions, with the most frequently observed values (modes) 690 corresponding to 0.235 m.day<sup>-1</sup>, 0.017 for  $K_{sat}$  and  $\mu$ , respectively. The same soil texture 691 could be found at the Arrou agricultural catchment (located 50 km south of Chartres, France), 692 where the measured value of  $K_{sat}$  at subsoil depth equals 0.41 m.day<sup>-1</sup> (Zimmer et al., 1995). 693 Therefore, the published data values of  $K_{sat}$  and  $\mu$  are close and consistent with the estimated 694 values obtained by the VDA method. 695

Furthermore, Jamagne et al. (1977) and Goulet et al. (2004) sought to determine the water 696 holding capacity (WHC) function of certain soil parameters (e.g. water content at field 697 capacity and wilting point, apparent density). In these studies, the WHC ranged between 1 698 mm.cm<sup>-1</sup> and 1.75 mm.cm<sup>-1</sup> for silty soil. Based on the drain depth of the Chantemerle field 699 (90 cm), these values correspond to 90 mm and 157.5 mm, respectively. The estimates using 700 VDA are therefore consistent with these values and lie within the WHC variation interval. 701 The estimated SIDRA-RU parameters are listed in Table 7 alongside the field-scale observed 702 data. 703

# **Table 7.** Estimated SIDRA-RU parameters using VDA compared with those found in the

705

literature

Source	$K_{sat}$ (m.day <sup>-1</sup> )	μ(-)	S <sub>max</sub> (mm)
Variational calibration	0.19; 0.23	0.044	112; 121
(Lagacherie, 1987)	0.235	0.017	-
(Bouarfa, 1995)	0.4	0.026	
(Zimmer et al., 1995)	0.41	-	-

Jamagne et al. (1977) &	-	-	[90, 157]
Goulet et al. (2004)			

One can see that VDA performs slightly better in terms of metrics evaluation (e.g. NSE, 706 KGE) than the PAP-GR calibration. These results indicate that VDA could be seen as a 707 relevant methodology to improve the predictive performance of the SIDRA-RU model. While 708 the adjoint model of SIDRA-RU has been successfully obtained through the automatic 709 differentiation (AD) tool TAPENADE, other large-scale models may need substantial 710 711 programming effort and forward model adaptation to get the adjoint code. As an alternative, some methods, including the reduced adjoint approach, have been proposed to avoid 712 laborious adjoint implementations (Antoulas, 2005; Altaf et al., 2013) but they require many 713 function evaluations and converge towards approximate solutions (LeGresley and Alonso, 714 2000). Overall, the task of generating the adjoint of any upgraded version of the conceptual 715 SIDRA-RU model looks straightforward, given that certain coding rules are being respected. 716

#### 717

#### 4.3. Water balance simulation

A recent drainage modeling study, which aims to demonstrate the benefit of using a multi-718 objective minimization approach to improve the estimation of HYDRUS parameters, has 719 shown a significant deviation (25%) between the cumulative measured and simulated water 720 volumes, which corresponds to 50 mm in just one drainage season (Turunen et al., 2020). In 721 722 this paper, a relevant result confirms the strong ability of VDA to simulate accurately the volume of water exported from drained agricultural fields. In fact, less than 1% ratio between 723 measured and simulated total drained water volume is obtained using VDA ( $\Delta V = 4.8 \text{ mm}$ 724 over five years of data, less than 1 mm/year). The robustness of the VDA method 725 implemented with the SIDRA-RU model has been successfully validated in Section 3.2.1 726 using the split sample test. It allows for a robust estimation of model parameters consistent 727 with the *in-situ* values found in the literature. Moreover, the estimation of SIDRA-RU 728

parameters using VDA instead of the PAP-GR calibration approach improves the daily discharge simulations and the cumulative water volume predictions, in particular. Finally, it would be worthwhile to investigate the impact of the sampling interval on the water volume prediction by considering weekly or be-weekly discharge observations instead of daily ones (Oubanas et al., 2018).

#### 734 **4.4. Local minima issues**

Local minima issues is commonly encountered when solving ill-posed problems in variational formulations involving a gradient-based minimization. The application of sophisticated global optimization methods (e.g. Monte Carlo) could be useful to deal with strongly nonlinear systems (Hoteit, 2008). Here, we investigate a preliminary ensemble minimization framework, aiming to reach a stable global minima solution.

Consequently, the convergence of  $K_{sat}$  and  $\mu$  parameters towards a global minimum during 740 741 the single optimization experiment highlights the strong consistency of the physically based SIDRA-RU module. However, the convergence pattern for parameters SSDI and Sinter 742 differs: dependent on the priors, the minimization process converges to different cost function 743 values, which underscores the presence of local minima or equifinality issues. In fact, these 744 two parameters manage the conceptual part of the SIDRA-RU model; their contribution is 745 746 limited to a short period, essentially during the first few days of each drainage season. Our results suggest that their values may require a dynamic year update during the minimization 747 process. Additional information on the reservoir capacity or beginning of the drainage season 748 749 may be needed to constrain further the minimization protocol concerning these reservoir parameters. 750

It is worth noting that the convergence speed may depend on the modeling approach (e.g.conceptual or physical, distributed or not), the choice of the optimization algorithm, and the

42

parameter type. In fact, SIDRA module parameters (e.g.  $K_{sat}$ ,  $\mu$ ) find a stable solution following 20 iterations while conceptual ones (e.g. *SSDI*,  $S_{inter}$ ) generally need 30 iterations. This explains the notable difference observed between the cost function thresholds.

Chossat and Saugnac (1985) highlight a potential relationship between  $K_{sat}$  and  $\mu$  for 756 different soil textures. Our finding suggests that these two groups of parameters are 757 complementary and may compensate each other, which typically leads to the occurrence of 758 local minima and equifinality issues (nearly the same prediction obtained with different 759 parameter values). The equifinality issue can be mitigated when assimilating additional 760 information by the VDA, e.g. water level variations h(t). We suggest herein that 761 complementary information (namely nitrate concentrations [NO3-] and fluxes) will 762 additionally constrain the solution in the upgraded SIDRA-RU version. The featured methods 763 have been implemented in order to improve both the understanding and estimation of the 764 SIDRA-RU input parameters. The present work can be extended to other drainage models. 765

#### 766 **4.5. The potential of the hybrid BV approach**

The implementation of the hybrid BV framework improves the stability of the obtained 767 solution compared to the classical VDA. That is, the estimated parameters are more 768 consistent from one study period to the other, which demonstrates the robustness of the 769 proposed method. In fact, the estimated parameter values on P1 and P2 are similar (Table 6), 770 in particular for the parameters  $K_{sat}$  and  $S_{max}$  (the reported difference between P1 and P2 771 estimated values are 0.027 mm.day<sup>-1</sup> and 5.3 mm, respectively). Moreover, the approach 772 offers an automatic tool to select, cleverly and objectively, the most suitable and robust set of 773 parameters, which was not the case by using visual interpretations (section 3.3). It is 774 worthwhile in this context to discuss the results of the latest SIDRA-RU developments 775 (Jeantet et al., 2021; Henine et al., 2022), which highlight the performance of the model in 776

terms of discharge simulations and the robustness of the estimated parameters. It was 777 observed that conceptual based parameters (e.g.  $S_{max} = SSDI + S_{inter}$ ) show a large deviation 778 range [50 mm, 150 mm]. These estimated values remain scattered, regardless of their 779 consistency as compared to the literature-based values (c.f. Table 7, last row). Otherwise, the 780 limited number of information available in the literature makes complex any confrontation 781 between the estimated conceptual parameters of the model and the field measurement values. 782 783 Moreover, the structure of the conceptual block of the model may need some additional information to be able to manage the related parameters in a robust way and with a reduced 784 785 level of uncertainty (e.g. hourly discharge observations, time-variant estimation of the conceptual parameters). In addition, let us note that any reliable conclusions on the relative 786 performance of the two methods (hybrid BV against variational DA) cannot be made on a 787 single test example, even within the identical twin experiment framework, where the 'true' 788 values of parameters are known. This is significantly more difficult when considering 789 realistic data and an approximate model (which is often the case with conceptual models). 790

#### 791 **5. CONCLUSIONS**

Our first objective has been to evaluate the impact of the SIDRA-RU input parameters on the model output, to rank them and form a sufficient control set. This has been achieved using both local and global adjoint sensitivity analysis. Then, the parameters from the control set have been estimated by assimilating the drainage discharge data, collected from the 36-ha Chantemerle agricultural field for the period 2008-2013, using VDA.

In addition to the local adjoint SA, the global adjoint SA method has been investigated in application to the SIDRA-RU model. Based on estimating the upper bound of the "totaleffect" Sobol indices, the GASA method has shown results consistent with the classical Sobol method (applied to the same model). In particular, the results of this analysis demonstrate that the two most influential parameters are  $K_{sat}$  and  $\mu$ . The GASA method has proven to be capable of estimating the impact of all input parameters at a very low computational time ( $\approx$ 2.6 s). As a result, only 4 (out of 11) parameters have been selected for the calibration step.

804 The VDA method has been successfully implemented on the conceptual-analytical model SIDRA-RU. This method has been validated using the 'split sample test' and compared to 805 Michel's step-by-step calibration method (PAP-GR). Slightly better performance could be 806 noted with the VDA method, which allowed estimating the best parameter set and 807 reproducing a satisfactory daily drainage discharge simulation (KGE = 0.76 for PAP-GR vs. 808 0.78 for VDA). An important advantage of these estimates has been revealed by comparing 809 the total cumulative drainage volumes. The parameters estimated using the VDA method 810 811 yielded a cumulative volume close to the observed volume at the field outlet (not exceeding 4.8 mm and less than 1% discrepancy) compared to values obtained with PAP-GR (58.7 mm, 812 for a 10% error). Moreover, the corresponding parameter values are consistent with those 813 found in the literature under similar soil conditions (i.e. texture, climate). A planned 814 forthcoming work will be to assimilate the nitrate concentration observations using the 815 upgraded version of SIDRA-RU, which will integrate a newly developed nitrate transport 816 model. We expect that the advantages of the implemented approach will be particularly 817 beneficial for the coupled drainage/nitrate model, since it requires yearly estimations of a 818 819 given input variable.

Lastly, the local minima issue has been explored. Our results here are in line with findings reported in the literature. It has been confirmed that using gradient-based minimization may lead to multiple possible solutions due to local minima. Therefore, a novel hybrid BV method has been implemented. Starting from an ensemble of optimal solutions given by VDA, the posterior PDF is computed using a specially defined likelihood function. At least, in theory, the hybrid BV method has to provide a more robust estimate of the parameters (in average sense, i.e. over a representative set of the test cases). The fact that significantly different estimates for *SSDI* and  $S_{inter}$  have been obtained in the split sample test underlines the importance for further investigation in this direction.

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#### 834 DATA AVAILABILITY

Bata archiving is underway. An upload within OZCAR network is scheduled (ORACLE
https://gisoracle.inrae.fr/).

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#### 1088 APPENDIX A

#### 1089 A.1. calculation of the corrected evapotranspiration (CET)

	Case 1: If $(S(t) > 0.6 * S_{inter})$	$CET = \beta * PET$	
1090	Case 2: If $(S(t) < 0.6 * S_{inter})$	$CET = \beta * PET * exp\left(-\frac{0.6*S_{inter} - S(t)}{S(t)}\right)$	(A.1)
	Case 3: If $(S(t) < 0)$	CET = 0	

#### 1091 A.2. Recharge (R) and water level in the soil reservoir (S) calculation

 $\begin{array}{ll} \mbox{Case 1: If } (0 < S(t) < S_{inter} \ ) & S(t) = S(t-1) + (P-CET); & R(t) = 0 \\ \mbox{Case 2: If } (S_{inter} < S(t) < S_{max} \ ) & S(t) = S(t-1) + (1-\alpha) * (P-CET); & R(t) = \alpha * (P-CET) \\ \mbox{Case 3: If } (S(t) < 0 \ ) & S(t) = 0; & R(t) = \alpha * (P-CET) \\ \end{array}$ 

#### 1093 A.3. Water table level and discharge calculation

h(t) = h(t - 1) + 
$$\left(R(t) - K_{sat} \frac{h(t-1)^2}{l^2}\right) / \mu C$$
  
1094 Q(t) = AJ(h(t)) + (1 - A)R(t)

$$Q(t) = AJ(h(t)) + (1 - A)R(t)$$

$$J(h(t)) = K_{sat} \frac{h(t)^2}{l^2}$$
(A.3)