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U-CARE: Utilities for performing goodness of fit tests and manipulating CApture– REcapture data

Rémi CHOQUET, Jean-Dominique LEBRETON, Olivier GIMENEZ, Anne-Marie REBOULET and Roger PRADEL.

7 Corresponding author : R. Choquet (remi.choquet@cefe.cnrs.fr), Centre d'Ecologie Fonctionnelle et Evolutive - CNRS - UMR 5175 - 1919 Route de Mende - F34293 Montpellier cedex 5 Abstract: In any statistical analysis, assessing the goodness of fit of a model to the data is crucial to avoid drawing incorrect conclusions. U-CARE is a computer application that deals with the mandatory first steps of the analyses of capture-recapture data: the preparation of the data set and the assessment of the fit of a general model (Cormack-Jolly-Seber and variants for single-state data; Jolly-Move and variants for multi-state data). U-CARE implements the current state of the art in goodness-of-fit testing by incorporating components aimed at detecting the most likely departures from assumptions (Pradel et al. 2003), (Pradel et al. 2005). It is a free and stand-alone application for Windows.

Keywords: Arnason-Schwarz model, Contingency tables, Directional tests, Memory model, Over dispersion coefficient, Test of multinomial mixture, Test WBWA, Test 3G.SR, Test M.ITEC, Transient,
 Trap-dependence.

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The estimation and comparison of demographic rates in animal and plant populations (Lebreton et al. 1992), (Gregg and Kery 2006) is currently based on following marked individuals over time and possibly space and analyzing the resulting data with adequate statistical models. Such models, in which individuals may move among states (such as geographical sites or breeding statuses) over discrete time periods, must account for the non-exhaustive detection of individuals. These so-called capture-recapture models are still evolving (Pradel et al. 2005). In their current form, they are based on two key assumptions.

- 57 1) The individuals are assumed to be independent,
- 582) Within the history of each individual, the successive pairs of each release associated with its59subsequent reencounter (if any) are assumed to be independent (Burnham 1991).

60 While fitting multi-state capture-recapture models can now be carried out without much difficulty 61 with a variety of computer software applications (M-SURGE (Choquet et al. 2004 #340), E-SURGE 62 (Rouan et al., Choquet et al. 2008) or MARK (White and Burnham 1999) for a frequentist approach, 63 MARK for Bayesian inference), the examination of the above mentioned assumptions can be done in 64 an optimal way (Pradel et al. 2003) only with program U-CARE. And yet this step is critical. For 65 instance, if a major structural effect such as the presence of transient individuals on the study site is 66 overlooked, survival will be severely underestimated. More generally, spurious effects will often be 67 detected if no model fits the data. This is best seen when considering how model selection proceeds. 68 In the frequentist paradigm, model selection is generally based on the Akaike information criterion 69 (AIC, see (Burnham and Anderson 2002)). The AIC is equal to the deviance plus two times the number 70 of estimable parameters. In presence of lack of fit, the deviance tends to be inflated, thus leading to 71 the selection of over-parameterized models and potentially to erroneous biological conclusions. 72 Moreover, the estimate of precision of the Maximum Likelihood Estimators (MLE) will be overly 73 optimistic. The consequences of lack-of-fit are thus too deleterious to be ignored.

74 A preliminary assessment of Goodness–Of–Fit (GOF) is thus a crucial prerequisite, just as in any 75 statistical analysis (D'Agostino and Stephens 1986). Unlike in regression, residuals are not easily 76 available in CR to check the validity of a model. Furthermore, the omnibus approach to goodness-of-77 fit testing that consists of comparing expected vs. observed sufficient statistics is impractical due to 78 the sparseness of the data. A specific approach was developed (reviewed in Pradel et al. 2003). By 79 making this approach available in an easy-to-use software application, we aim at encouraging 80 practitioners to assess the validity of their assumptions. Sound model selection, i.e. preceded by 81 appropriate GOF assessment, is indeed becoming more and more common in the literature (e.g. 82 (Henaux et al. 2007), (Jenouvrier et al. 2008)) but it is still far from being systematic. We describe 83 here the main features of program U-CARE. More details on tests and tools can be found in the user's 84 manual (Choquet et al. 2005).

U-CARE (<http://purl.oclc.org/NET/U-CARE>) is a free, easy-to-use, stand-alone, menu-driven 85 computer application for Windows, with a set of options for managing data and GOF tests 86 87 capabilities for both single-state and multi-state capture-recapture models (see Figure 1). The two 88 kinds of test are presented as the sum of components examining different aspects of the data 89 through a range of contingency tables. This structure guides the choice of an appropriate model. For 90 instance, if the only significant subcomponent of the multistate goodness-of-fit test is TEST 3G.SR, a 91 two age-class structure for survival should be used (Pradel et al. 1995), while, if this is TEST 92 WhereBeforeWhereAfter (WBWA), a memory model is recommended (Pradel et al. 2003). Due to 93 historical reasons, biological relevance is particularly furthered for single-state data through the 94 computation of directional tests. We briefly examine hereafter the following four aspects of data 95 analysis for which U-CARE can be very useful:

- 96 1) How to prepare the data?
- 97 2) How to detect overdispersion?
- 98 3) How to correct for overdispersion?
- 99 4) How to conduct specific tests?

TRANSFORM DATA GOODNESS-OF-FIT	GOODNESS-OF-FIT for Multistate TOOLS						
Clean data	TEST WBWA						
Suppress first encounter	TEST 3G.SR						
Select occasions	TEST 3G.Sm						
Gather occasions	TEST M.ITEC						
Reverse time	TEST M.LTEC						
Multistate to single state	Sum of Multistate TESTs over groups						
Recode states	OPTIONS •						
Encounter histories layout	TOOLS						
Individuals layout	M-ARRAY						
Extract groups	P-Value						
Suppress groups	Test of mixtures for contingency table						
Pool groups	Test of homogeneity for contingency table						
Split for trap-dep analysis	Power divergence family statistics used for contingency table test •						
Build age-groups for trap-dep analysis							
Build groups for trap analysis (Fletcher)							
Sorting cohorts into BIOMECO files							

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Figure 1 : U-CARE has 5 menus showing utilities for input-output of capture-recapture files (FILE menu), manipulation of
 data (TRANSFORM DATA menu), goodness-of-fit testing for single-state (GOODNESS OF FIT menu) and multi-state
 (GOODNESS OF FIT FOR MULTI-STATE menu) data and various additional tools (TOOLS menu). Three of them are shown.

104 1) <u>How to prepare the data?</u>

First, the menu TRANSFORM DATA offers some tools for selecting a data subset, such as a particular range of years or some groups (e.g. males or females, 'ringed as young' or 'ringed as adults'), and for recoding data, such as pooling states, groups or years. Among other features, the menu FILE makes it possible to convert data between two main formats (Biomeco (Lebreton and Roux 1989) and Mark (White and Burnham 1999)).

110 2) How to detect overdispersion?

111 When the data are ready, the next step should be the assessment of the fit of a general model to the 112 data at hand. For single-state data, the classical model is the model with time dependent only 113 parameters (Cormack-Jolly-Seber) (CJS); for multistate data, this is a time and state dependent model 114 (Arnason-Schwarz) (AS). For both of them, the probability of the encounter history of each individual is calculated conditional on its first capture. Optimal goodness of fit tests of the assumptions
inherent in the CJS or the JMV model, this last being a slight generalization of the AS model, are then
derived based on the classical partitioning, according to sufficient statistics T, of the likelihood
P(data/parameters):

119 P(data/parameters)=P(data/T) x P(T/parameters).

120 The compatibility of the data with the hypergeometric distributions in P(data/T) are tested 121 asymptotically by contingency table chi-squared tests. These tests are organized into several 122 interpretable components by further partitioning P(data/T) (Pradel et al. 2005). If systematic 123 departures from the general model can be ruled out, lack of fit may result from the lack of independence among individuals (the two members of a pair, social groups, a contagious disease...). 124 125 Then, the value of the Pearson statistic is much larger than the residual degrees of freedom. One simple measure for overdispersion is \hat{c} defined as the ratio of the Pearson statistic X² by its number 126 of degrees of freedom df 127

$$\hat{c} = \frac{\chi X^2}{df}$$

This ratio can be calculated for each component individually or overall. Although there is no clear cut 129 130 decision rule to decide that an observed lack of fit results solely from overdispersion, a reasonable rule of thumb is: the \hat{c} ratio is greater than 1 for all components, and there is no component for 131 which the \hat{c} ratio greatly exceeds the others. This procedure, although not perfect, works generally 132 well as long as the overall ratio does not exceed 3, sometimes 5 (Burnham and Anderson 2002 -133 134 numéro de page?); if it does, an important factor has likely been left out of the model and should be identified. On the other hand, small structural effects that result in an overall \hat{c} between 1 and 3 can 135 136 without much damage be assimilated to overdispersion and treated as noise (see below for how to 137 do this in practice).

To illustrate these general ideas, we will now treat the example of the study of movements of Canada geese *Branta canadensis* between three wintering regions, Mid-Atlantic, Chesapeake and Carolinas, between 1984 and 1989 (Hestbeck et al. 1991).

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The option 'GOODNESS-OF-FIT for Multi-state' of the main menu of U-CARE opens onto the 142 goodness-of-fit test of the JollyMove (JMV) model (Brownie et al. 1993). Multi-state models 143 144 (Hestbeck et al. 1991) allow for transitions between states (the wintering sites for the geese), 145 survival probabilities and encounter probabilities. In the JMV model, transitions vary by state of 146 departure, state of arrival and time interval. Survival probabilities vary by state of departure and time interval. Encounter probabilities vary by previous state, current state and date. This model, in 147 contrast with the better-known Arnason-Schwarz (AS) model, allows encounter probabilities to vary 148 149 by previous state. However, there is currently no optimal GOF test available for the AS model. 150 Program U-CARE contains specific tests for transience (Test 3G.SR, null hypothesis H0: 'There is no 151 difference in the probability of being later reencountered between "new" and "old" individuals 152 encountered simultaneously.'), trap-dependence (Test M.ITEC, H0(i): 'There is no difference in the 153 probabilities of being reencountered in the different states at i+1 between the animals in the same 154 state at occasion i whether encountered or not encountered at this date, conditional on presence at 155 both occasions.'), and memory (Test WBWA H0: 'There is no difference in the expected state of next reencounter among individuals previously encountered in the different states'). The \hat{c} ratio for the 156 JMV model is computed from these three main components plus two complementary tests 3G.SM 157 158 and M.LTEC (Pradel et al. 2005).

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$$\hat{c} = \frac{\chi X_{WBWA}^2 + \chi X_{3G,SR}^2 + X \chi_{3G,SM}^2 + \chi X_{M,ITEC}^2 + X \chi_{M,LTEC}^2}{df_{WBWA} + df_{3G,SR} + df_{3G,SM} + df_{M,ITEC} + df_{M,LTEC}}$$

160 If we were to consider that the lack of fit is due solely to overdispersion, the overall \hat{c} ratio for the 161 geese would be:

$$\hat{c} = \frac{472.8 + 117.7 + 302.7 + 68.2 + 21.9}{20 + 12 + 119 + 27 + 19} = 5$$

- 163 However, the \hat{c} ratio relative to the WBWA component is huge: 23.6 (see Figure 2)!
- 164

TEST WB	WA (tes	t of me	mory: W	here Be	fore vs	Where	After)	,	group	=1
======	=== Sumi	mary of	TEST W	BWA res	ults ===	======	==			
sta, pv	al, df,	occ, s	tate, 1	if fis	her O ot	herwis	se			
19.591	0.000	2.000	2.000	1.000	0.000					
37.868	0.000	2.000	2.000	2.000	0.000					
4.487	0.034	1.000	2.000	3.000	1.000					
80.590	0.000	1.000	3.000	1.000	0.000					
98.761	0.000	4.000	3.000	2.000	0.000					
0.807	0.369	1.000	3.000	3.000	1.000					
27.705	0.000	1.000	4.000	1.000	0.000					
53.694	0.000	2.000	4.000	2.000	0.000					
25.293	0.000	1.000	4.000	3.000	0.000					
43.655	0.000	1.000	5.000	1.000	0.000					
50.926	0.000	2.000	5.000	2.000	0.000					
29.476	0.000	2.000	5.000	3.000	0.000					
	TES	T WBWA	(sta, p	val, df)					
472.855	0.000	20.000								

Figure 2 Details of Test WBWA for the Canada goose (*Branta Canadensis*) data from U-CARE. The Pearson statistic (*sta*) is
 calculated per occasion and state. The *c* ratio is 23.6. This suggests strong memory effects in accordance with the overall
 p-value (*pval*).

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170 3) How to correct for overdispersion?

171 Given the large value of the overall \hat{c} ratio obtained for the Canada geese, we must try and \hat{c} find a 172 more general model that takes into account the effects unveiled by the GOF test: \hat{c} \hat{c} the strong 173 memory effect and a transient effect. The \hat{c} ratio of Test 3G.SR is indeed 9.8. A model studied in 174 (Rouan et al.) combines these two effects. After discarding the components corresponding to the 175 effects incorporated in this model (memory: component WBWA; transience: component 3G.SR), the 176 new \hat{c} ratio is

$$\hat{c} = \frac{302.7 + 68.2 + 21.9}{119 + 27 + 19} = 2.4,$$

This value is typical of large data sets with several thousands of individuals where individual 178 179 differences inherent in any animal population are inevitably detected. This ratio can be used as a variance inflation factor for the model of Rouan and derived models. There is an option in M-SURGE, 180 E-SURGE and MARK to introduce this factor. The impact of the correction factor \hat{c} is null on 181 parameter estimates, but estimated (co)variances are multiplied by \hat{c} , widths of the Wald-CI-182 intervals are $\sqrt{\hat{c}}$ times larger and the deviance is divided by \hat{c} . This in turn affects model selection 183 where the AIC is replaced with the QAIC (Burnham and Anderson 2002). This procedure of identifying 184 a starting model is valid more generally and should be pursued until an acceptable value for the \hat{c} 185 ratio is reached. 186

187 4) How to conduct more specific tests?

U-CARE provides details that may give some additional information, notably the tables of observed
 and expected numbers. For example, Figure 3 shows that newly marked individuals are less
 recaptured than previously marked individuals at occasion 5 on sites 1 and 3 but not 2. This is a
 general pattern over all occasions which suggests that transients may be absent from the central site
 A more specific model could be considered. For an application to seasonal data, see ((Gauthier et

al. 2001)). U-CARE also gives directional tests (REF) focusing more closely on the detection oftransience and trap-dependence.

```
TABLE 3G.SR(5,1): animals caught at occasion5 in state1 seen earlier and later
     Seen again - Never seen again
New
     193 317
Old
    279
          602
Table of expected values
     Seen again - Never seen again
                    291.859
New
     1118.141
    1253.859
                   627.141
old
TABLE 3G.SR(5,2): animals caught at occasion5 in state2 seen earlier and later
     Seen again - Never seen again
    297
New
           748
01d |512 1330
Table of expected values
     Seen again - Never seen again
New
     292.8317
                    752.1683
old
    516.16834
                     1325.8317
TABLE 3G.SR(5,3): animals caught at occasion5 in state3 seen earlier and later.
     Seen again - Never seen again
     1157
New
          504
    |134 197
old
Table of expected values
     Seen again - Never seen again
     193.9022
                    467.0978
New
     |97.09778
                    233.9022
Old
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196 Figure 3 Tables of expected and observed values for Test 3G.SR at occasion 5 for sites 1, 2, 3

197 Eventually, the option 'TOOLS' contains some specific tools, not easily found, like the ability to test 198 for mixtures of multinomial in a contingency table (Pradel et al. 2003), with an improved algorithm.

To cite U-CARE or acknowledge its use, cite this Software Note as follows, substituting the version of the application that you used for "Version 2.3":

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