

Evidence of a large carnivore population recovery: Counting bears in Greece

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1	Evidence of a large carnivore population recovery: Counting bears in
2	Greece
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- 12 Short title: Bear abundance in Greece

1 Abstract

- 2
- 3 **Keywords**: brown bear, Greece, Jolly-Seber capture-mark-recapture modelling
- 4 framework, nonivasive genetic sampling, population monitoring, power pole hair
- 5 traps, Ursus arctos
- 6

1 Introduction

2 The brown bear (Ursus arctos) is a charismatic key species of the biodiversity of 3 the Holarctic region (Servheen, Herrero & Peyton, 1999), capable of attracting 4 attention and resources to conservation efforts through its function as an iconic 5 flagship species (Simberloff, 1999). Once, extirpated and threatened by 6 extinction throughout large parts of Europe (Zedrosser et al., 2001), bears have 7 made a remarkable recovery, with populations increasing and expanding in 8 several countries (Deinet et al., 2013). The benefits for nature conservation 9 notwithstanding, such recoveries create the potential of increased wildlife 10 mortality and human - wildlife conflicts (Gardner et al., 2010). In this context, 11 informed, science-based conservation and management decisions are urgently 12 required to safeguard the recovery and survival of small brown bear populations, 13 which has been identified as an important European conservation priority 14 (Habitat Directive 92/43/CEE). Since conservation priorities often arise in relation 15 to bear numbers, reliable population assessments are fundamental in 16 conservation and management planning (Mowat & Strobeck, 2000, Lorenzini et 17 al., 2004, Kendall et al., 2008).

Brown bears in Greece reach their southern-most distribution in Europe and are therefore an important aspect of European biodiversity. They belong to the Dinaric-Pindos (DP) bear population, which has been identified as one of the largest and most important on the European continent (Zedrosser *et al.*, 2001); the DP bear population appears to be stable, numbering more than 3,000 individuals (Deinet *et al.*, 2013, Kaczensky *et al.*, 2013). In Greece, brown bears are considered to be endangered (Mertzanis, Giannakopoulos & Pylidis, 2009).

1 In recent years the increase of extra-limital appearances of bears throughout the 2 country (Karamanlidis, Krambokoukis & Kantiros, 2008), and the increase in 3 negative human – bear interactions (i.e., mainly damages by bears to human 4 property) (Karamanlidis et al., 2011) and bear – vehicle collisions (Karamanlidis 5 et al., 2012) suggest that the brown bear population in Greece might be recovering. The arising conservation issues of this population expansion and 6 7 growth make the development and implementation of an effective Action Plan for 8 the species an immediate conservation priority (Mertzanis et al., 2009).

9 Despite public interest and costly management programs, no rigorous and 10 reliable estimates of abundance and density of brown bears in Greece exist. 11 Because of their crepuscular and nocturnal activity patterns, large home range 12 size, shyness, low density and the rugged and forested terrain they inhabit, counting brown bears by traditional field methods (e.g., direct observations, 13 14 extrapolations of counts of mother with cubs of the year) is expensive or 15 sometimes even unfeasible (Mace et al., 1994). Traditional field methods are 16 viewed as less rigorous than more efficient current capture-mark-recapture 17 (CMR) and DNA-based methods (Garshelis, 2006). The size given therefore for 18 the brown bear population in Greece [i.e., 190 – 260 individuals (Mertzanis et al., 19 2009)], which is presumably a combination of expert knowledge and observation 20 counts, should be considered an educated guess that might not reflect actual 21 population processes (Nichols & Williams, 2006). The current lack of reliable data 22 for the brown bear population in Greece sharply contradicts with the need to 23 develop a sound, effective and outreaching management and conservation 24 strategy in the country (Ciucci & Boitani, 2008).

1 Genetic methods play a pivotal role in bear conservation and cannot be 2 neglected when developing effective conservation strategies and planning the 3 long-term survival of threatened bear populations (Swenson, Taberlet & 4 Bellemain, 2011). One of the most attractive applications of molecular methods, 5 commonly used now by management agencies (McCall et al., 2013), is the possibility of estimating the (minimum) number of free-ranging individuals 6 7 (Schwartz, Tallmon & Luikart, 1998; Kohn et al., 1999). Coupled with noninvasive 8 genetic sampling (NGS), which is ideal for studying small, endangered 9 populations as it precludes the unnecessary capture of individuals, and CMR 10 methods, genetic study methods have been successfully deployed to obtain bear 11 abundance and density estimates for large (Mowat & Strobeck, 2000, Boersen, 12 Clark & King, 2003, McCall et al., 2013) and small study areas with irregular 13 sampling designs and small population sizes [<100 individuals (Gardner, Royle & 14 Wegan, 2009, Obbard, Howe & Kyle, 2010, Gervasi et al., 2012, Latham et al., 15 2012)].

The goal of our study was to apply noninvasive genetic sampling as part of a capture–mark–recapture (CMR) study design for estimating bear abundance and density in Greece.

19 Methods

20 Study area

Brown bears in Greece occur in two disjunct populations in the western and eastern part of the country (Figure 1A); the population in the western part of the country is stable or locally increasing and is estimated to number approximately 190 – 230 individuals (Mertzanis *et al.*, 2009). The study was carried out in the

1 western part of the brown bear range in Greece, at the approximately 250 km-2 long Greek part of the Pindos mountain range, extending from the Albanian and 3 the Former Yugoslav Republic (FYR) of Macedonian - Greek border in the 4 northwest of the country to the region of Central Greece (Figure 1B). The study 5 area forms a mosaic of elevations and habitats, from deep canyons and low altitudes to steep and high mountains. Elevations range from a couple of hundred 6 7 meters above sea up to the peak of Mount Smolikas at 2,637 m above sea level. 8 This wide range of elevations results in two major forest zones: A higher 9 elevation conifer zone, with tree species such as the endemic Greek fir (Abies 10 cephalonica), the European black pine (*Pinus nigra*) and the Bosnian pine (*Pinus* 11 heldreichii) and a lower elevation mixed broadleaf forest, which is dominated by 12 oak (Quercus sp.), beech (Fagus sp.) and fir (Abies sp.). The local fauna 13 includes other large and medium-sized carnivores, such as the grey wolf (Canis 14 lupus) and the golden jackal (Canis aureus), and intensity of human activity 15 ranges from areas with low human presence to major cities with populations > 16 50,000.

17 Sampling methods

Genetic sampling was based on the natural marking and rubbing behavior of bears (Green & Mattson, 2003) on wooden poles of the telephone and electricity network (hereafter power poles) in Greece (Karamanlidis *et al.*, 2007) and therefore we did not use any attractant to lure bears to the sampling stations. In the preparatory phase of the study (2002-2006) we inspected and evaluated the suitability of power poles as non-invasive genetic sampling stations according to predefined criteria that considered among others marking and rubbing frequency

and intensity, type of habitat, vegetation density and intensity of human activity (Karamanlidis, 2008); then we conducted a pilot study that confirmed the suitability of power poles for the non-invasive genetic monitoring of brown bears (Karamanlidis *et al.*, 2010).

5 For the purposes of this study and following the inspection and evaluation of 4147 power poles in the study area we selected 171 and placed barbed-wire hair 6 7 traps (Kendall & McKelvey, 2008) on them, thus creating four different sampling 8 areas (i.e., Vitsi - Varnoundas, and Northern, Central and Southern Pindos) 9 (Figure 1B, Table 1). We selected location and size of the sampling areas so as 10 to cover the maximum of the core range of the species in the Pindos mountain 11 range; sampling areas were separated either by distance or/and geo-12 morphological features (e.g. extensive plains and valleys) and human 13 infrastructure (e.g. major highways). We calculated the size of the sampling 14 areas by drawing a buffer zone of 5.86 km around the Minimum Convex Polygon 15 (MCP) defined by the outermost sampling stations in each study area; 5.86 km 16 was the mean distance between all the individual bear recaptures recorded 17 during the study. We carried out systematic sampling efforts in all four sampling 18 areas from October 2007 to October 2010, when poles were inspected and hair 19 samples were collected monthly.

20 Figure 1

Figure 1 A) Map of southeastern Europe. The shaded areas indicate the approximate geographic range of brown bears in the region. B) Geophysical map of Greece indicating the location of the study area. C) Map of Greece indicating the location of four sampling areas for the non-invasive collection of brown bear

- 1 genetic samples (2007–2010). D) The inset photograph pictures an adult female
- 2 brown bear with her cubs rubbing and marking a power pole.
- 3 Table 1 Descriptive statistics of four sampling areas for the non-invasive
- 4 collection of genetic samples from brown bears in Greece (2007–2010).

Sampling area	Location	Sampling area size (km²)	Sampling stations (N)
Vitsi – Varnoundas	N: 40.6935, E: 21.3192	1270.39	40
Northern Pindos	N: 40.3707, E: 21.0907	1066.22	45
Central Pindos	N: 39.9197, E: 21.4054	999.72	47
Southern Pindos	N: 39.6305, E: 21.3711	954.33	39
Total		4290.66	171

5 Genetic Methods

6 Each tuft of hairs on a set of barbs was considered a sample; hair samples were 7 collected without contact to human skin, were placed in uniquely numbered 8 paper envelopes and then stored at room temperature in zip-lock bags with silica 9 gel (Roon, Waits & Kendall, 2003) until being analyzed by Wildlife Genetics 10 International (Nelson, British Columbia, Canada), a laboratory that specializes in 11 non-invasively obtained samples that often have low quantities of DNA. DNA was 12 extracted using the DNeasy Blood & Tissue kits (QIAGEN, Hilden, Germany), 13 following the manufacturer's instructions. All extractions took place in a facility in 14 which amplified DNA had never been handled before.

To determine individual identity each sample was genotyped at the microsatellite loci G1D, G10J, G10L (Paetkau & Strobeck, 1994, Paetkau, Shields & Strobeck, 1998), G10C, G10P (Paetkau *et al.*, 1995), MU51 and MU59

(Taberlet *et al.*, 1997). Gender identification was established through the analysis
of the amelogenin gene (Ennis & Gallagher, 1994). Up to 10 additional loci were
analyzed for ≥1 sample from each individual to enable more detailed population
genetic analyses. These extended genotypes were used to confirm differences
between individuals with similar 7-locus genotypes.

6 Thermal cycling was performed using a MJ Research PTC100 thermocycler 7 with 96 well 'Gold' blocks (MJ Research Inc., St. Bruno, Quebec, Canada). 8 Polymerase chain reaction (PCR) buffers and conditions were used according to 9 Paetkau et al. (1998), except that markers were not co-amplified, because co-10 amplification may reduce the success rates for hair samples (D. Paetkau, Wildlife 11 Genetics International, personal communication). Two mM MgCl2 was used for 12 all markers except G10J (1.8 mM). An automated sequencer (ABI 310) was used, 13 and genotypes were determined using ABI Genescan and Genotyper version 2.1 14 software (Perkin Elmer-Applied Biosystems, Foster City, California, USA). The 15 sizing of the PCR products was performed using capillary electrophoresis.

16 Mixed samples (samples with hair from >1 bear) were reliably identified by 17 evidence of ≥ 3 alleles at ≥ 1 locus (Roon *et al.*, 2005). To minimize genotyping 18 errors in the final data set, low-quality and putatively mixed samples were 19 excluded from further analyses (Paetkau 2003). Genotypes were replicated for all 20 1) individuals identified in one sample, 2) pairs of individuals that differed at only 21 1 or 2 loci (1- and 2-mismatch pairs) and 3) pairs of individuals that differed at 3 22 loci when ≥ 1 locus was consistent with allelic dropout. Test for allelic dropout, 23 presence of null alleles, and scoring errors caused by stutter peaks were

performed with Micro-Checker version 2.2.3 (van Oosterhout, Hutchinson & Wills,
 2004).

3 We used the observed number of alleles (A) and expected heterozygosity $(H_{\rm e})$ to express genetic variation in our population. We calculated the probability 4 5 of identity (P_{ID}) and of siblings (P_{SIB}) to describe the power of our markers to identify individuals (Waits, Luikart & Taberlet, 2001) using the software GIMLET 6 7 version 1.3.2 (Valière 2002). To allow for the possibility of mismatches caused by 8 genotyping error, we also looked for the pairs of genotypes that were matched at 9 all but 1, 2, and 3 loci (1-MM, 2-MM, and 3-MM pairs) using the program GenAIEx 6 (Peakall and Smouse 2006). 10

11 Estimating Abundance and Density

12 To estimate abundance, we used the Jolly-Seber (JS) CMR modeling framework 13 that allows accounting for detectability less than one and dealing with open 14 populations, i.e., systems in which births and deaths occur (Jolly, 1965, Seber, 15 1965). JS models provide estimates of abundance, probabilities of survival, 16 detection and recruitment [sensu (Schwarz & Arnason, 1996) i.e., the probability 17 that an animal from the hypothetical super-population would enter the population 18 on a particular time interval. We analyzed each of the four regions separately. 19 We considered time and sex effects as possibly explaining environmental and 20 individual variation in these demographic rates, entering in an additive or 21 interactive way in the model, which led to a total of 125 candidate models. To 22 select among these models, we used the Akaike's Information Criterion corrected 23 for small sample size (AICc) and considered the model with the lowest AICc 24 value as being the model that best balanced bias and precision (Burnham &

1 Anderson, 2002). In addition, we used changes in AICc values ($\Delta AICc$) to 2 compare model support with reference to the model best supported by the data. 3 Annual abundance estimates were obtained by model averaging in which each 4 model contributed to the final estimate according to its AICc weight (Burnham & 5 Anderson, 2002). The analyses were performed using POPAN (Arnason & Schwarz, 1995) available in MARK (White & Burnham, 1999) that was called 6 7 from R (R Development Core Team, 2011) with package RMark (Laake & 8 Rexstad, 2008).

9 We tested the validity of important assumptions underlying the safe use of the 10 JS model like the presence of transient individuals or an effect of trapping using 11 standard goodness-of-fit tests (Pradel, Gimenez & Lebreton, 2005) as 12 implemented in program U-CARE (Choquet *et al.*, 2009).

13 Although we considered a potential effect of sex on the detection probability, 14 other sources of individual heterogeneity might go undetected in the present 15 analysis, which could lead to severe bias in abundance estimates (e.g., 16 (Cubaynes et al., 2010b). This might be for example due to individuals being 17 more active than others at the proximity of power poles, hence making them 18 more capturable. To assess whether some individual heterogeneity was present 19 besides gender effects, we compared for each of the four study areas the model 20 best supported by the data to its counterpart, incorporating heterogeneity in the 21 detection probability. To do so, we considered finite-mixture JS models (Pledger, 22 Pollock & Norris, 2010) that extend standard JS models by assuming that the 23 animals come from different classes of detection, although we do not know which 24 class each individual is from. Heterogeneous JS models were fitted using the R

1 package 'hetage' that is available from http://homepages.ecs.vuw.ac.nz/~shirley/. 2 To compare homogeneous and heterogeneous models, AIC could not be used 3 because of the violation of regularity conditions (e.g. (Gimenez & Choquet, 2010) 4 in a CMR context), and we resorted to a likelihood ratio test that was distributed 5 under the null hypothesis of homogeneous detection probabilities as a 50:50 6 mixture of chi-square distributions with 0 and 1 degrees of freedom (Self & Liang, 7 1987). Note that, to get the P-value of this test, using the mixture corresponds to 8 halving the P-value from using the standard chi-square distribution with 1 degree 9 of freedom.

We calculated bear density using the sampled area size of each sampling network respectively (i.e., Vitsi – Varnounds: 1270.39 km²; Northern Pindos: 1066.22 km²; Central Pindos: 999.72 km²; Southern Pindos: 954.33 km², Table 1). Less than 3% of the study areas were not suitable habitat for bears (e.g., lakes, rivers and human settlements); therefore we retained these areas in the sampled area and density calculations.

16 **Results**

17 Sampling Effort

From 2007-2010, we conducted 4422 inspections to the 171 sampling sites and
collected 860 hair samples (Table 2). We collected bear hair at 92.25% of all hair
traps (i.e., Vitsi – Varnoundas: 95%; Northern Pindos: 100%; Central Pindos:
95%; Southern Pindos: 79%) during 25.75% of all power pole visits (i.e., Vitsi –
Varnoundas: 30%; Northern Pindos: 34%; Central Pindos: 26%; Southern
Pindos: 13%).

24 Genotypic Success, Marker Power, and Quality Control

1	Of the 860 samples sent to the laboratory, we excluded 250 (29%) from the
2	analysis due to insufficient genetic material (i.e., hair without follicles). Of the 610
3	samples that we attempted to analyze, 11 (1.8%) appeared to contain DNA from
4	>1 individual, and 154 samples (25%) failed the DNA extraction process. From
5	the remaining 445 samples (64% of total collected) 211 (61F, 151M) individuals
6	were identified and 102 individuals were recaptured (Table 2). From the 211
7	individuals identified, 206 individuals (98%) produced complete 7-locus
8	genotypes plus gender assignment.

9 Table 2. Descriptive statistics of efforts at four sampling areas in Greece (2007 -

10	2010) for the	non-invasive	collection	of brown	bear hair	samples
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Sampling area	Visits	Bear presence	Hair samples collected	Females identified	Males identified	Total <i>N</i> identified	<i>N</i> of animals recaptured
Vitsi – Varnoundas	1115	358	215	13	39	52	28
Northern Pindos	1146	421	295	26	51	77	29
Central Pindos	1220	351	249	19	42	61	33
Southern Pindos	941	132	101	3	18	21	12
Total	4422	1262	860	61	150	211	102

11 Mean observed heterozygocity across the 7 markers used to identify 12 individuals was 0.70 (Table 3). The probability that 2 randomly drawn, unrelated 13 individuals would share the same genotype (P_{ID}) was 0.0000005, and the 14 probability that full siblings would have identical genotypes (P_{SIB}) was 0.002 15 (Table 3). Based on the observed distribution of genotype similarity for the 7 loci 16 used for individual identifications, we predicted that no pair of matching 17 genotypes could exist within our dataset. Of the 211 individuals present in our

- 1 analysis, 91% had \geq 10-locus genotypes and, when all available loci were
- 2 considered, all individual bears differed at \geq 3 loci.
- 3 **Table 3.** Variability of microsatellite markers used to determine individual identity

Marker	<i>H_E^a</i>	H _o ª	A ^a	P_{ID}^{a}	P _{SIB} ^a
G1D	0.78	0.80	6	0.08	0.37
MU59	0.78	0.76	8	0.07	0.38
G10C	0.76	0.76	6	0.09	0.39
G10P	0.75	0.67	9	0.10	0.40
G10J	0.70	0.77	6	0.14	0.43
G10L	0.57	0.55	5	0.23	0.52
MU51	0.55	0.57	5	0.24	0.53
X	0.70	0.70	6.42		
Overall probability of identity				5.299e-07	0.002

4 of brown bears in Greece, 2007-2010.

^a H_E = expected heterozygosity; H_O = observed heterozygosity; A = number of alleles; P_{ID} = probability of identity; P_{SIB} = probability of identity among siblings

5 Abundance and density estimation

6 The standard assumptions of the JS model were valid as showed by the goodness-of-fit tests that were non-significant for all regions ($\chi_3^2 = 2.45$, p-value = 7 0.49 for Vitsi – Varnoundas; $\chi_4^2 = 3.58$, p-value = 0.47 for Northern Pindos; $\chi_7^2 =$ 8 0.47, p-value = 0.99 for Central Pindos; $\chi_2^2 = 0.71$, p-value = 0.70 for Southern 9 Pindos). Model selection showed contrasted effects between regions (Table 4). 10 11 Survival was constant in Vitsi - Varnoundas and Southern Pindos and sex-12 specific in Central and Northern Pindos with the addition of time in the latter. 13 Recruitment was constant in all regions but Northern Pindos in which it was sex 1 and time dependent on an additive scale. Detection was constant in all regions 2 but Vitsi – Varnoundas in which it was sex-specific. There was substantial 3 uncertainty regarding the best model to select for all regions, with several models 4 within the range $\Delta AICc < 2$, which motivated the use of model-averaged 5 abundance estimates (see below). We did not detect extra heterogeneity in the detection probability ($\chi_1^2 = 2.5$, p-value = 0.06 for Vitsi – Varrnoundas; $\chi_4^2 = 2.6$, 6 p-value = 0.31 for Northern Pindos; χ_7^2 = 8.5, p-value = 0.15 for Central Pindos; 7 8 $\chi_1^2 = 0.6$, p-value = 0.22 for Southern Pindos).

9 Table 4. Model selection results from CMR analysis of the brown bear 10 populations in Greece using the Jolly-Seber modelling framework. For each of 11 the four regions (i.e., Vitsi - Varnoundas, Northern, Central and Southern Pindos), 12 the Akaike's Information Criterion value corrected for small sample sizes (AICc), 13 the difference in AICc value between the *i*th model and the model with the lowest 14 AICc value (Δ AICc). Constant (.), sex (s) and time (t) effects were considered on 15 survival (φ), detection (p) and entry (b) probabilities either as main effect, in an additive (+) or in interactive (x) fashion. Note that, for each of the four regions, 16 17 only the 10 top ranked models are displayed although 125 models were originally 18 fitted to the data.

Region	Model	k	AICc	ΔAICc
	φ(.) p(s) b(.)	5	95.17	0.00
	φ(s + t) p(.) b(s)	7	96.49	1.34
	φ(s) p(.) b(s)	6	97.51	2.36
Vitsi-	φ(t) p(s) b(.)	6	97.51	2.37
varnoundas	φ(.) p(s) b(s)	6	97.54	2.39
	φ(.) p(s) b(t)	6	97.57	2.42
	φ(s) p(s) b(.)	6	97.58	2.43

	φ(s) p(.) b(.)	5	97.91	2.76
	φ(s + t) p(.) b(.)	6	98.26	3.12
	φ(.) p(.) b(s)	5	98.37	3.22
	$\phi(s + t) p(.) b(s + t)$	7	126.03	0.00
	$\phi(s) p(.) b(s + t)$	8	126.19	0.16
	$\phi(s + t) p(s) b(s + t)$	8	127.97	1.94
	$\phi(s) p(t) b(s + t)$	11	128.02	2.00
Northorn Dindos	$\phi(s) p(s) b(s + t)$	10	128.47	2.48
Northern Findos	$\phi(s + t) p(.) b(t)$	11	128.67	2.64
	φ(s) p(.) b(t)	8	128.82	2.79
	φ(s) p(t) b(t)	9	129.09	3.06
	φ(s) p(t) b(s)	9	129.41	3.38
	$\varphi(s + t) p(t) b(s + t)$	9	129.95	3.92
	φ(s) p(.) b(.)	8	143.03	0.00
	φ(s) p(.) b(t)	8	143.64	0.61
	φ(.) p(s) b(.)	9	143.77	0.74
	φ(s + t) p(.) b(.)	7	144.56	1.54
Control Dindoo	φ(.) p(.) b(.)	10	144.73	1.71
Central Findos	φ(s) p(.) b(s)	9	144.75	1.72
	φ(s + t) p(.) b(t)	7	144.76	1.73
	φ(s) p(.) b(s + t)	10	144.82	1.79
	φ(.) p(s) b(t)	8	144.82	1.79
	φ(s) p(s) b(.)	11	144.83	1.80
	φ(.) p(.) b(.)	4	55.05	0.00
	φ(s) p(.) b(.)	5	55.14	0.09
	φ(.) p(s) b(.)	6	57.84	2.78
	φ(t) p(.) b(.)	7	57.96	2.91
Southorn Dindoc	φ(.) p(.) b(s)	5	57.98	2.93
Southern Findos	φ(s) p(s) b(.)	6	58.00	2.95
	φ(.) p(.) b(t)	6	58.04	2.99
	φ(s + t) p(.) b(.)	5	58.28	3.23
	φ(s) p(.) b(s)	8	58.39	3.34
	φ(s) p(.) b(t)	6	58.41	3.36

The most abundant region was Vitsi – Varnoundas with 68 individuals on
 average over the study period, followed by Northern Pindos with 53 individuals,
 Central Pindos with 51 individuals and Southern Pindos with 10 individuals
 (Table 5).

Table 5. Abundance estimates from CMR analysis of the brown bear populations in Greece using the Jolly-Seber modeling framework. For each of the four regions (i.e., Vitsi - Varnoundas and Northern, Central and Southern Pindos), the time and sex-specific population size estimate, as well as the lower (LCL) and upper (UCL) confidence limits of the 95% confidence interval obtained from model averaging are given.

Region	Sex	Estimate			
		_	2007-8	2008-9	2009-10
		value	18.7	22.4	21.4
	female	LCL	4.9	7.1	5.2
		UCL	71.1	70.8	88.3
		value	43.0	47.4	50.8
Vitsi - Varnoundas	male	LCL	12.4	15.1	15.6
		UCL	148.7	148.6 10 69.8 7	165.2
		value	61.7	69.8	72.2
	pooled	LCL	19.7	24.7	23.6
		UCL	193.0	197.6	220.6
		value	16.9	19.7	4.0
	female	LCL	8.7	10.1	0.7
		UCL	32.8	007-82008-918.722.44.97.171.170.843.047.412.415.148.7148.651.769.819.724.793.0197.616.919.78.710.132.838.544.946.520.626.797.781.351.866.331.138.7	24.5
Northern Pindos		value	44.9	46.5	27.7
Northern Findos	male	LCL	20.6	26.7	9.1
		UCL	97.7	81.3	84.6
	noolod	value	61.8	66.3	31.8
	pooled	LCL	31.1	38.7	9.9

		UCL	122.8	113.6	102.1
		value	22.2	14.5	11.4
	female	LCL	12.3	6.5	3.5
		UCL	40.1	32.4	37.2
		value	41.9	34.2	28.7
Central Pindos	male	LCL	26.4	21.2	13.9
		UCL	66.6	55.0	58.9
		value	64.1	48.6	40.1
	pooled	LCL	39.1	28.9	18.1
		UCL	105.0	81.9	88.6
		value	2.1	1.0	0.9
	female	LCL	0.7	0.2	0.2
		UCL	6.0	4.8	3.8
		value	11.6	8.2	7.0
Southern Pindos	male	LCL	7.6	5.4	3.8
		UCL	17.5	12.7	12.1
		value	13.6	9.3	7.7
	pooled	LCL	8.3	5.5	4.1
		UCL	22.3	15.6	14.4

1 Discussion

We present the first rigorous estimate of brown bear abundance and density in the western range of brown bears in Greece. By using genetic data from hair samples collected from poles of the electricity and telephone network, an approach that has not been implemented in this scale before, we obtained precise population estimates despite fairly low capture and recapture rates. The remarkable consistency of our estimates between years (especially the years 2008 and 2009) lends credibility to our results.

9 There are several assumptions that need to be checked to use the Jolly-Seber 10 model and obtain reliable abundance estimates. First, we assumed that there

1 was no immigration or emigration (i.e., closure assumption) in the four regions 2 we studied. This was not entirely true as several individuals moved from region 3 XX to region XX during the study period. A modelling option would have been to 4 use multisite JS models (e.g., (Lebreton et al., 2009) instead of considering the 5 region as a factor, hence allowing the estimation of movement between regions. However, with so few individuals actually making the transitions between regions, 6 7 fitting multisite JS models would be numerically unstable with movement 8 parameters estimated on the boundary (Gimenez et al., 2005) and identifiability 9 issues (Gimenez, Choquet & Lebreton, 2003). Add here a sentence on how we 10 should interpret our abundance estimates consequently. A second assumption 11 underlying the application of JS models is that all animals, marked or unmarked, 12 should have the same capturability (i.e., homogeneous capturability assumption). 13 Here, there was no reason to believe that the use of poles to "capture" individuals 14 would distinguish between marked and unmarked individuals. Third, the duration 15 of the sampling period should be short in comparison to the time between 16 sampling occasions. In our study, most of the observations were made over eight 17 months, which was considered a short enough with regard to the yearly time 18 interval we used between sampling periods as bear survival was high (Hargrove 19 & Borland, 1994). Fourth, and probably the most crucial assumption, detection 20 probabilities were assumed homogeneous between individuals, heterogeneity if 21 present and ignored being source of strong bias in the abundance estimates 22 (Cubaynes et al., 2010a). We found a sex effect for the Vitsi - Varnoundas 23 population, males being more detectable than females (0.18 [SE=0.12] for 24 females and 0.38 [SE=0.08] for males), and this difference was accounted for 25 when calculating abundance estimates. We also tested for extra individual

heterogeneity not explained by differences between males and females, but did not find any, which suggest that the risk of bias in abundance estimates was avoided. Last but not least, it is usually assumed that individuals retain their mark throughout the study and that these marks are read correctly. When interpreted in the context of DNA marking, these two assumptions refer to data quality and the possibility of genotyping error. Last but not least, issue of the coverage of the poles.

8 Management implications

9 Importance of estimating abundance for conservation and management.
10 Originality of the poles as non-invasive monitoring method. Very efficient when
11 combined with capture-recapture methods that account for detection < 1 and
12 demography (mortality / birth).

13 Conclusions

14 Acknowledgements

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