

Impact of Missing Data on Mixtures and Clustering

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Impact of Missing Data on Mixtures and Clustering

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Take home message

Missing data may change preconceptions

- Mixtures: EM has unexpected behaviour concerning degeneracy dynamic
- Clustering: the missing data pattern may convey some information on partition

These topics are in the research agenda of Statisticians since:

The larger the datasets, the more missing data may appear. . .

- 1 Introduction
- 2 Impact of MAR data on the EM algorithm
- 3 Impact of MNAR data on clustering
- 4 Concluding remarks

1 Introduction

Introduction

Missing data: notations

- $y = \{y_1 | \dots | y_n\}^T$: full dataset with n individuals
- $\mathbf{y}_i = (y_{i1}, \dots, y_{id}) \in \mathcal{Y} = \mathbb{R}^d$ with individual $i \in \{1, \dots, n\}$
- $\mathbf{c} = \{c_1 | \dots | c_n\}^T \in \{0,1\}^{n \times d}$: pattern of missing data for the full dataset
- $c_i = (c_{i1}, \dots, c_{id}) \in \{0, 1\}^d$: pattern of missing data for individual $i \in \{1, \dots, n\}$

$$c_{ij} = 1 \Leftrightarrow y_{ij}$$
 is missing

- y_i^{obs} : the observed variables values for indiv. i (and $y^{\text{obs}} = \{y_1^{\text{obs}} | \dots | y_n^{\text{obs}} \}^T$)
- y_i^{mis} : the missing variables values for individual i (and $y^{\text{mis}} = \{y_1^{\text{mis}}|\dots|y_n^{\text{mis}}\}^T$)

Typology of the missingness mechanisms

■ Missing completely at random (MCAR):

$$p(c|y;\psi) = p(c;\psi) \quad \forall y$$

■ Missing at random (MAR):

$$p(c|y; \psi) = p(c|y^{obs}; \psi) \quad \forall y^{mis}$$

■ Missing not at random (MNAR): the mechanism is not MCAR nor MAR

A missing mechanism is ignorable if likelihoods can be decomposed as

$$L(\theta, \psi; \ \underline{y^{\rm obs}, c} \) = L(\psi; c|y^{\rm obs}) \times L(\theta; y^{\rm obs})$$

Some simple algebra show that this occurs when missing mechanism is not MNAR

Inference of θ

"If the missing mechanism is ignorable then likelihood-based inferences for θ from $L(\theta; y^{\text{obs}})$ will be the same as likelihood based inference for θ from $L(\theta, \psi; y^{\text{obs}}, c)$." [Little and Rubin, 2002 Section 6.2]

- M(C)AR is ignorable
- MNAR is not ignorable

- Partition (*K* clusters): $z = (z_1 | \dots | z_n)^T \in \{0,1\}^{n \times K}$ where

 - **u** $z_i = (z_{i1}, \dots, z_{iK}) \in \{0, 1\}^K$ **u** $z_{ik} = 1$ if y_i belongs to cluster k, $z_{ik} = 0$ otherwise
- Mixture model: $y_1, ..., y_n$ are i.i.d. from the d-variate Gaussian mixture

$$f(y_i; \pi, \theta) = \sum_{k=1}^K \pi_k f_k(y_i; \theta_k)$$

- $\pi_k = p(z_{ik} = 1), \ \pi = (\pi_1, \dots, \pi_K)$
- $f_k(.;\theta_k) = \phi(.;\mu_k,\Sigma_k)$ is the d-variate **Gaussian distribution** with mean vector μ_k and covariance matrix Σ_ν
- $\theta = (\theta_1, \dots, \theta_K)$ is the whole mixture parameter
- Clustering: MAP principle from the mixture output to estimate the partition

Outline

- 2 Impact of MAR data on the EM algorithm
 - Gaussian mixture degeneracy without missing data
 - Gaussian mixture degeneracy with missing data

Degeneracy genesis: unbounded likelihood

d-variate Gaussian mixture

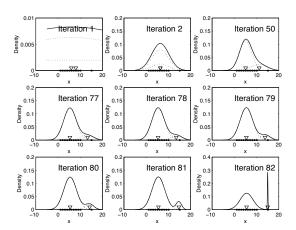
$$f(y_i; \boldsymbol{\theta}) = \sum_{k=1}^{K} \pi_k \underbrace{\frac{1}{(2\pi)^{d/2} |\Sigma_k|^{1/2}} \exp\left(-\frac{1}{2} (y_i - \mu_k)^T \Sigma_k^{-1} (y_i - \mu_k)\right)}_{\phi(y_i; \mu_k, \Sigma_k)}$$

- Sampling: $v_1, \ldots, v_n \overset{i.i.d.}{\sim} p(:;\theta)$ without any missing data $(v^{\text{obs}} = v)$
- Likelihood: $\ell(\theta; y) = \ln L(\theta; y) = \sum_{i=1}^{n} \ln f(y_i; \theta)$

particular center
$$\mu_2 = y_i \quad \Rightarrow \quad \lim_{|\Sigma_2| \to 0} \ell(\theta; y) = +\infty$$

[Kiefer and Wolfowitz, 1956] [Dav. 1969]

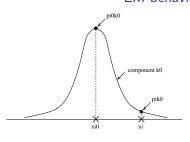
EM behaviour: illustration



- degeneracy may occur even when starting from large variances
- convergence can be slow when far from the degenerate limit
- convergence extremely fast near degeneracy



EM behaviour: results



$$u_0 = \left[\frac{1}{p_{i_0k_0}}, \{p_{ik_0}\}_{i \neq i_0}\right]$$

degeneracy of component k_0 at y_{i_0}

$$||u_0|| \to 0$$

[Biernacki and Chrétien, 2003] [Ingrassia and Rocci, 2009]

Proposition 1: Existence of a bassin of attraction

 $\exists \epsilon > 0 \text{ s.t. if } \|u_0\| \le \epsilon \text{ then } \|u_0^+\| = o\|u_0\| \text{ with probability } 1.$

Proposition 2: Speed towards degeneracy is exponential

 $\exists \epsilon>0,\, \alpha>0 \text{ and } \beta>0 \text{ s.t. if } \|u_0\| \leq \epsilon \text{ then, with probability } 1, \\ |\Sigma_{k_0}^+| \leq \alpha/|\Sigma_{k_0}| \cdot \exp\Big(-\beta/|\Sigma_{k_0}|\Big).$

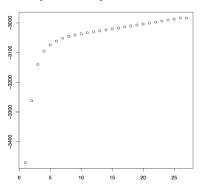
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 - Gaussian mixture degeneracy with missing data

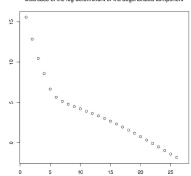
EM behaviour illustration

- Breast cancer tissue of the UCI database repository: 106 units, 9 variables.
- 10% of missing data randomly generated
- K = 4 clusters

Log-likelihood according to the number of iterations



Decrease of the log-determinant of the degenerated component



Detail from the illustration

							7		
	1	2	3	4	5	6	- /	8	9
1	211.00		0.09	30.75	151.98	4.94	14.27	27.24	217.13
2	196.86	0.02	0.09	28.59	82.06	2.87	7.97	27.66	200.75
3	144.00	0.12	0.05	19.65	70.43	3.58		7.57	160.37
4	172.52	0.13	0.04		192.22	5.12	19.32	32.19	174.93
5	121.00	0.17	0.09	24.44	144.47	5.91	22.02	10.59	141.77
6	223.00	0.12	0.08	33.10	197.01	5.95	30.45	12.96	252.48
7		0.17	0.23	34.22	94.35	2.76	31.28	13.88	180.61
8	303.00	0.06	0.04	22.57		4.54	21.83	5.72	321.65
9	250.00	0.09	0.09	29.64	180.76	6.10	26.14	13.96	280.12
10	391.00	0.06	0.01	35.78		7.41	22.13	28.11	400.99
11	176.00	0.09	0.08	20.59	79.71		18.23	9.58	191.99
12	145.00		0.11	21.22	82.46	3.89	20.30	6.17	162.51
13	124.13	0.13	0.11	20.59			18.46	9.12	134.89
14	103.00	0.16	0.29	23.75	78.26	3.29	22.32	8.12	124.98

Table: Data belonging to the degenerated component.

- Cvg. towards a degenerated component (no plateau of the log-likelihood)
- Degeneracy relatively slow: log-likelihood linear according to the nb of it.
- Number of points of the degenerated solution greater than the space dimension
 d (but the number of complete points lower than d)

Intermediate conclusion on missing data

Like the complete data v case

- Likelihood is unbounded
- EM can be attracted by degenerate solutions

Unlike the complete data y case

- Risk to consider a degenerated solution as valid
- Risk of losing a lot of time in useless iterations

Statisticians should be aware of such dangerous EM behaviour...

... since missing data are more and more frequent

Understanding degeneracy speed on a toy example

- Univariate framework, no mixture, only one observed data: y
- Maximum likelihood estimator (Unbounded likelihood!): $\mu = y$, $\sigma^2 = 0$
- Suppose equivalently that n-1 data are unobserved (unchanged likelihood)
- \blacksquare Here is one iteration of a (useless) EM algorithm (it. q)

$$\mu^{(q+1)} = \frac{(n-1)\mu^{(q)} + y}{n} \quad \text{and} \quad \sigma^{2(q+1)} = \frac{(n-1)\sigma^{2(q)} + (y - \mu^{(q+1)})^2}{n}$$

Linear grow of the log-likelihood (have a look also when n increases!)

$$\ell(\theta^{(q)}; x) \sim -0.5q \log \frac{n-1}{n}$$

Geometrical convergence rate towards 0 for the variance

$$\sigma^{2(q)} \sim \sigma^{2(0)} \left(\frac{n-1}{n}\right)^q$$

Influence of the missing data rate

% missing data	0	5	10	15	20	25	30
% deg.	16	4	12	11	46	51	100
Average nb of iterations before deg.	2	13	13	82	304	138	215

Table: Frequency and speed of degeneracy (deg.) according to the rate of missing data on the breast cancer data set

When the rate of missing data increases:

- The rate of degeneracy increases
- The number of iterations before degeneracy seems to (globally) increase

Again, statisticians should be aware of such dangerous EM behaviour. . .

... since missing data are more and more frequent

Impact of MNAR data on clustering

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Outline

- 3 Impact of MNAR data on clustering
 - A model-based MNAR clustering approach
 - Medical study illustration

Proposed MNAR models in clustering

Question we address now

Since MNAR is not ignorable, which distribution $p(c|y, z; \psi)$ to propose?

Hypothesis 1: conditional independence

$$p(c_i|y_i, z_{ik} = 1; \psi) = \prod_{j=1}^d p(c_{ij}|y_i, z_{ik} = 1; \psi)$$

Hypothesis 2: linear function within canonical link functions ρ

$$p(c_{ij} = 1 | y_i, z_{ik} = 1; \psi) = \rho(\alpha_{kj} + \beta_{kj}y_{ij})$$

- $\begin{aligned} & \quad \quad \psi = (\alpha,\beta) \text{ where } \alpha = (\alpha_{11},\ldots,\alpha_{1d},\ldots,\alpha_{K1},\ldots,\alpha_{Kd})^T \in \mathbb{R}^{Kd} \text{ and } \\ & \quad \quad \beta = (\beta_{11},\ldots,\beta_{1d},\ldots,\beta_{K1},\ldots,\beta_{Kd})^T \in \mathbb{R}^{Kd} \end{aligned}$
- lacksquare ρ is the cdf of any continuous distribution (logit, probit)

	Effect variabl		Effect on members	the class hip <i>k</i>	Nb parameters	
	Depends on <i>j</i>	Depends on k	Depends on j	Depends on k	Continuous	
MNARz ^j y ^k	✓	✓	✓	✓	2Kd	
MNAR <i>yz^j</i>	✓	×	✓	✓	(K+1)d	
MNARy ^k z	✓	✓	Х	✓	K(d+1)	
MNARyz	✓	×	X	✓	(K+d)	
MNARy	✓	×	X	×	d	
$MNARy^k$	✓	✓	Х	×	Kd	
MNARz	X	×	×	✓	K	
$MNAR z^j$	×	×	✓	✓	Kd	

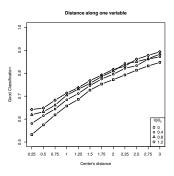
Remarks:

- MNAR $z^j y^k$ is the most complex model
- MNARz, MNARz^j: the only effect of missingness is on the class membership, $\psi = (\alpha_{11}, \dots, \alpha_{1d}, \dots, \alpha_{K1}, \dots, \alpha_{Kd})^T$, $\mathsf{p}(c_{ij} = 1 \mid y_i, z_{ik} = 1; \psi) = \rho(\alpha_{ki})$
- MCAR is a specific and simple case

MNARz analysis: pattern c gives information on partition z!

Draw Bayes error of a MNARz model with two components and 20% of missing data

$$\pi_k = 0.5$$
, $\|\mu_2 - \mu_1\|$ varies, $\Sigma_1 = \Sigma_2 = I$, $|\alpha_2 - \alpha_1|$ varies



Both μ_k and α_k act on the Bayes error

Reinterpretation of the MNARz and MNARz^j models as MAR

Commonly used in Machine Learning [Jones, 1996], [Little and Rubin, 2002], [Josse et al., 2019]

Mixture model for y^{obs} and Bernoulli distribution for $C \Leftrightarrow \mathsf{MAR}$ mixture model for $\tilde{y}^{\mathrm{obs}} = (y^{\mathrm{obs}}|c)$

For example,

$$y^{\text{obs}} = \begin{pmatrix} ? & 2.6 & 5 \\ \text{blue} & 1.9 & 4 \\ \text{red} & 2.3 & ? \end{pmatrix}, \quad c = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 0 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

then \tilde{y}^{obs} is expressed as

$$\tilde{y}^{\mathrm{obs}} = \left(\begin{array}{ccccc} ? & 2.6 & 5 & 1 & 0 & 0 \\ \mathsf{blue} & 1.9 & 4 & 0 & 0 & 0 \\ \mathsf{red} & 2.3 & ? & 0 & 0 & 1 \end{array} \right).$$

Proposition 3: in terms of maximum likelihood

The maximum likelihood estimate associated to the dataset \tilde{y}^{obs} under MAR model is the one associated to the dataset y^{obs} under MNARz or MNARz^j models.

Identifiability

Previous works: [Teicher, 1963] (without NA), [Miao et al., 2016] (for MNAR data)

Proposition 4

Assume that

- **The marginal mixture** $\sum_{k=1}^{K} \pi_k f_k(y_i; \theta_k)$ is identifiable
- **2** There exists a total ordering \leq of $\mathcal{F}_i \times \mathcal{R}$, for $j \in \{1, ..., d\}$ fixed, where $\mathcal{F}_{i} = \{f_{1i}, \dots, f_{Ki}\}\$ and $\mathcal{R} = \{\rho_{1}, \dots, \rho_{K}\} = \{\rho(.; \psi_{1}), \dots, \rho(.; \psi_{K})\}.$ The total ordering is s.t. $\forall k < \ell$, $F_k = \rho_k f_{ki} \leq F_\ell = \rho_\ell f_{\ell i}$ implies

$$\lim_{u\to+\infty}\frac{\rho_{\ell}(u)f_{\ell j}(u)}{\rho_{k}(u)f_{k j}(u)}=0$$

Then the mixture model with one of the MNAR* mechanisms is identifiable up to label swapping

All MNAR models are identifiable (or at least generically identifiable) for probit/logit

Estimation procedure overview

- Use EM or Stochastic EM (SEM) algorithms
- MNARz and MNARz^j: EM and SEM are very simple
- MNARy*: the SE step requires a within Gibbs loop, sometimes involving itself a Sampling Importance Resampling (SIR)

	E	M	SEM		
MNAR <i>z</i> MNAR <i>z^j</i>	,		✓		
	Probit	Logit	Probit	Logit	
MNARy*	no closed form	no closed form, optim. pb	✓	require algorithms as SIR (costly)	

Model selection

Can select between MCAR and MNAR* with any information criterion (BIC, ICL)

Even if the missing mechanism is ignorable for MCAR...

 \dots need to model c to compare a MCAR and a MNAR model

CAUTION

- It is just a selection between several proposed MNAR models
- It is not deciding if missingness procedure is "genererically" MNAR or not

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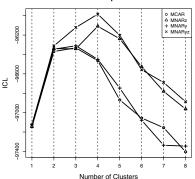
Hospital data description

- Number of patients: n = 5 146
- Number of features: d = 7
 - Age
 - Size
 - Weight
 - Cardiac frequency
 - Hemoglobin concentration
 - Temperature
 - Minimum Diastolic and Systolic Blood Pressure
- Percentage of missing data: 6.4%

Doctors are convinced that their missing data are MNAR

ICL comparison

ICL Comparison



- MCAR, MNARy and MNARz are equivalent until K = 3
- MNARz and MNARyz clearly indicate presence of an additional cluster (K = 4)

It seems to be an illustation of the effect of c through MNARz and MNARyz

4 Concluding remarks

To conclude

Summary

- Statisticians should properly consider the potential missing data impact both from an algorithmic and from a modeling point of view
- EM: be careful about degeneracy which seems to exacerbated/masqued
- MNAR: interest of the simple but meaningful model MNARz, link with usual methods

Ongoing works

- EM: propose mechanism to identify/discard degenerate runs
- MNAR: extend to categorical, count and mixed data

Thanks!