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Tokelo I Letshedi, Jean-Claude Malela-Majika, Philippe Castagliola, Sandile C Shongwe. Distribution-free triple EWMA control chart for monitoring the process location using the Wilcoxon rank-sum statistic with fast initial response feature. Quality and Reliability Engineering International, 2021, 37 (5), pp.1996-2013. 10.1002/qre.2842. hal-03504670

HAL Id: hal-03504670 https://hal.science/hal-03504670

Submitted on 29 Dec 2021

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Distribution-free triple EWMA control chart for monitoring the process location using the Wilcoxon rank-sum statistic with fast initial response feature

Tokelo I. Letshedi¹, Jean-Claude Malela-Majika¹, Philippe Castagliola², Sandile C. Shongwe¹

Abstract

The exponentially weighted moving average (EWMA) control chart is a memory-type chart known to be more efficient in detecting small and moderate shifts in the process parameter. The double EWMA (DEWMA) chart is an extension of the EWMA chart that is more effective than the latter in the detection of small-to-moderate shifts. This paper proposes a new distribution-free (or nonparametric) triple EWMA (TEWMA) control chart based on the Wilcoxon rank-sum (W) statistic to improve the detection ability in the process location parameter. Moreover, a new fast initial response (FIR) feature is added to further improve the sensitivity of the new TEWMA chart. The performance of the proposed TEWMA chart with and without FIR features is compared to those of the existing EWMA and DEWMA W charts. It is observed that the TEWMA chart with and without FIR features is superior to the competing charts in most situations. A real-life illustration is provided to show the application and implementation of the new chart.

Keywords: Double EWMA chart; EWMA chart; Fast initial response; Nonparametric chart; Overall performance; Run-length; Triple EWMA.

1. Introduction

All production lines and service-rendering establishments are recommended to use statistical process monitoring (SPM) techniques. This will assist with producing or rendering high quality services or products that will meet needs of end users with low risk benefits. In applications where very little information is known about the underlying process, it is mostly advised to use nonparametric or distribution-free charts rather than the parametric ones. This is because they do not depend on any specific underlying distribution and they are generally in-control (IC) robust; see for instance, the arguments in Chakraborti and Graham¹. For efficiency, memory-type control charts are usually used to detect small-to-moderate shifts of any process parameter. There are four different types of memory-type control charts in the SPM literature and these are: (i) Exponentially weighted moving average (EWMA) chart by Roberts², (ii) Cumulative sum (CUSUM) chart by Page³, (iii) Generally weighted moving average (GWMA) chart by Abbas⁵. For more detailed studies or literature reviews on the latter memory-type schemes, see Hawkins and Olwell⁶, Ruggeri et al⁷, Mabude et al⁸, Nawaz and Han⁹, Abid et al¹⁰, Adegoke et al¹¹ and Adeoti and Koleoso¹². The focus of this paper is

^{*}Corresponding author. E-mail: sandile@tuks.co.za;

¹Department of Statistics; College of Science, Engineering and Technology; University of South

Africa; Pretoria; South Africa.²Université de Nantes & LS2N UMR CNRS 6004, Nantes, France.

on improving the performance of the nonparametric EWMA-type charts to monitor the process location, especially during start-up period. Improving the performance of the EWMA-type control charts during start-up period has never been addressed for the nonparametric SPM context. For some recent contributions to nonparametric charts, refer for instance to Mukherjee et al¹³, Triantafyllou¹⁴, Chong et al^{15,16} and Shafqat et al¹⁷. Next though, a literature review of some related parametric research works are discussed and thereafter, the objective of this paper is outlined.

With an end goal to increase the responsiveness of the EWMA control charts to detect small shifts in a process mean, Shamma and Shamma¹⁸ proposed the double EWMA (DEWMA) control chart. Originally not aware of the latter paper, Zhang and Chen¹⁹ also studied the performance of the DEWMA chart to monitor the process mean; note though, there are additional aspects that were considered in the latter paper but absent in the former. Some additional studies on the DEWMA charts in different contexts can be found in Hong et al²⁰, Alkahtani²¹, Teh et al²², Chao et al²³ and, Alevizakos and Koukouvinos²⁴. Based on latter publications, it is generally accepted that the DEWMA charts are more sensitive than EWMA charts for small-to-moderate shifts. For the nonparametric case, Karakani et al²⁵ showed that the DEWMA based on the exceedance statistic is superior to the EWMA one in detecting small shifts.

The focus of this paper is on the Wilcoxon rank sum statistic, denoted as W. Note that Li et al^{26} presented the EWMA W chart with measurements utilizing a simple random sampling method. A while later, Malela-Majika and Rapoo²⁷ presented the EWMA W chart utilizing structured sampling techniques (i.e. based on rank set sampling method). More recently, Malela-Majika²⁸ proposed a DEWMA W chart which is shown to have better out-of-control OOC performance in detecting small-to-moderate shifts in the process location than the EWMA W chart. Other discussions on control charts based on the W statistic can be found in Tapang et al^{29} , Malela-Majika and Rapoo³⁰, Chakraborti and Graham^{1,31}, Triantafyllou³², Mabude et al^{33} , Tercero-Gomez et al^{34} and Qiu³⁵.

Therefore, the main objective of this paper is to improve further on the research works by Li et al^{26} and Malela-Majika²⁸, by proposing a new triple EWMA (TEWMA) *W* chart to monitor the process location shifts and then compare its performance to that of EWMA and DEWMA *W* charts. Note that there is no work in the SPM literature that addresses the use of TEWMA chart in a distribution-free context. To the best of authors' knowledge, Alevizakos et $al^{36,37}$ investigated the performance of the TEWMA chart for monitoring the process mean

and time between events, respectively. Moreover, Alevizakos et $al^{36,37}$ compared the performance of each TEWMA chart against the corresponding EWMA and DEWMA charts, where it is observed that the TEWMA charts have better performance in most situations. Note though, for thorough derivations of the run-length properties of the TEWMA chart for any independent and identically distributed (i.i.d.) process, readers are referred to Alevizakos et al^{36} .

The fast initial response (FIR), initially suggested in the SPM literature by Lucas and Crosier³⁷, is a head-start feature with similar (but, slightly more pronounced) positive effects compared to the time-varying control limits as it identifies issues during the start-up period. Thereafter, more researchers have used this enhancement technique at the start-up period of the monitoring process; see for instance Rhoads et al³⁹, Knoth⁴⁰ and Lu⁴¹. Note that Steiner⁴² proposed the use of a more formal multiplicative factor expression that yields tighter control limits at start-up and converges to the standard control limits as time increases. The basic FIR feature in Steiner⁴² is denoted as FIR in this paper. Later, Haq et al⁴³, proposed a modified FIR (MFIR) feature which is shown to yield better OOC performance than the Steiner⁴²'s FIR feature. Therefore, in addition to introducing the TEWMA *W* chart; in this paper, a new improved MFIR (IMFIR) feature is incorporated to the TEWMA *W* chart to improve its performance during the start-up period. For the sake of completeness and comparison purpose, the basic FIR and MFIR are also incorporated to the EWMA, DEWMA and TEWMA *W* charts.

The rest of the paper is structured as follows: Section 2 provides some design background of the W test statistic as well as the EWMA and DEWMA W charts. The design of the proposed TEWMA W charts with and without FIR features are discussed in Section 3. Section 4 gives an empirical performance of the proposed TEWMA W chart and some comparisons against the EWMA and DEWMA W charts. An illustrative example is given in Section 5 using real-life data. Finally, Section 6 provides some concluding remarks.

2. The design of the EWMA and DEWMA W charts

2.1 Wilcoxon rank sum test (W)

Assume that $X = \{x_i, i = 1, 2, ..., m\}$ is an IC reference (i.e. Phase I) sample with unknown cumulative distribution function (c.d.f.), F(x), and $Y_t = \{y_{tj}, t = 1, 2, 3, ...; j = 1, 2, 3, ..., n\}$, simply denoted as *Y*, is a set of test samples with c.d.f. G(y). The observations from the reference sample are assumed to be independent from each other and from the ones in the test

samples. However, the distribution of the test samples is the same as the one of the reference sample with a difference in the location shift only. In other words, $F(x) = G(x + \delta)$, where δ represents the change (or shift) in the location parameter. Thus, when the process is incontrol (IC), $\delta = 0$; hence, F = G. However, when $\delta \neq 0$, the process is said to be out-ofcontrol (OOC). Wilcoxon⁴⁴ proposed the W statistic by merging the m reference observations with the n observations of the t^{th} test sample together and re-arranging all values of the observations in (n+m)an ascending order. Let A = { $a_{(1)}, a_{(2)}, \dots, a_{(s)}, \dots, a_{(n+m-1)}, a_{(n+m)}$ } be the resulting ordered sample, where $a_{(s)} = x_l$ if x_l is the smallest l^{th} value from the newly formed sample A, also $a_{(s)} = y_k$ if y_k is the smallest k^{th} value from A, this means that $s \in \{1, 2, ..., n + m\}, k \in \{1, 2, ..., n\}$ and $l \in \{1, 2, ..., m\}.$

Let $I = \{a'_{(1)}, a'_{(2)}, \dots, a'_{(s)}, \dots, a'_{(m+n-1)}, a'_{(m+n)}\}$ be the sample of the indicators where $a'_{(s)} = 1$ if $a_{(s)} \in X$ and $a'_{(s)} = 0$ if $a_{(s)} \in Y$. Then, at time *t*, the *W* statistic is defined by

$$W_t = \sum_{s=1}^{n+m} (i \times a'_{(s)})$$
(1)

Assuming that there are no ties (i.e. duplicated values in the combined observations), the mean and the variance of the W statistic are respectively given by (see for instance Li et al²⁶):

$$E(W_t) = \mu_W = \frac{m(m+n+1)}{2}$$
(2)

and

$$Var(W_t) = \sigma_{W_t}^2 = \frac{mn(m+n+1)}{12}.$$

2.2 The EWMA W control chart

The EWMA W chart was proposed by Li et al^{26} and its charting statistic is defined by

$$E_t = \lambda W_t + (1 - \lambda) E_{t-1} \tag{3}$$

where λ ($0 < \lambda \le 1$) is a smoothing parameter and $E_0 = \mu_W$. After some basic algebraic manipulation, the mean and variance of (3) are respectively given by

$$E(E_t) = \mu_E = \mu_W$$

(4)

and

$$Var(E_t) = \sigma_{E_t}^2 = \sigma_W^2 \left(\frac{\lambda}{2-\lambda}\right) (1-(1-\lambda)^{2t}).$$

Hence, the upper and lower control limits of the EWMA W chart are computed as follows:

$$UCL_t/LCL_t = \mu_E \pm L_E \sigma_{E_t},\tag{5}$$

where $L_E > 0$ is the control limits coefficient which has to be calculated such that the attained IC *ARL* (*ARL*₀) is close or equal to the pre-specified nominal *ARL*₀. Thus, the EWMA *W* chart gives a signal if E_t plots beyond the control limits.

2.3 The DEWMA W chart

The DEWMA W chart was proposed by Malela-Majika²⁸ and its charting statistic is given by

$$D_t = \lambda E_t + (1 - \lambda)D_{t-1} \tag{6}$$

where E_t is as given in (3) and $D_0 = \mu_W$. That is, the smoothing procedure is applied twice, i.e. the first smoothing is done in (3) and the second one is done in (6). After some basic algebraic manipulation, the mean and variance of (6) are respectively given by

$$E(D_t) = \mu_D = \mu_W$$

(7)

and

$$Var(D_t) = \sigma_{D_t}^2 = \frac{\lambda(1 + \lambda'^2 - \lambda'^{2t}[(t+1)^2 - (2t^2 - 2t - 1)\lambda'^2 + t^2\lambda'^4])}{(1 + \lambda')^3} \sigma_W^2,$$

with $\lambda' = 1 - \lambda$. Hence, the upper and lower control limits of the DEWMA *W* chart are computed as follows:

$$UCL_t/LCL_t = \mu_D \pm L_D \sigma_{D_t} \tag{8}$$

where $L_D > 0$ is the control limits coefficient which has to be calculated such that the attained ARL_0 is close or equal to the pre-specified nominal ARL_0 . Thus, the DEWMA W chart produces an OOC signal when D_t value plots beyond the control limits in (8).

3. The design of the TEWMA W chart with and without FIR features

3.1 TEWMA W chart without FIR features

The TEWMA design entails applying the smoothing procedure three times; instead of twice as for the DEWMA or once as for the EWMA. Following on similar line of argument as Alevizakos et al^{36} , it follows that the charting statistic of the TEWMA *W* chart is given by

$$T_t = \lambda D_t + (1 - \lambda) T_{t-1}, \tag{9}$$

where D_t is given in (6) and $T_0 = \mu_W$. From (3), (6) and (9):

$$E_t = \lambda W_t + (1 - \lambda)E_{t-1}$$

$$D_t = \lambda E_t + (1 - \lambda)D_{t-1}$$

$$T_t = \lambda D_t + (1 - \lambda)T_{t-1},$$
(10)

that is, the original W_t statistic in (1) goes through a smoothing procedure three times; hence, simplifying each expression in (10) leads to the following:

$$E_{t} = \lambda \sum_{i=1}^{t} (1-\lambda)^{t-i} W_{i} + (1-\lambda)^{t} E_{0}$$

$$D_{t} = \lambda \sum_{i=1}^{t} (1-\lambda)^{t-i} E_{i} + (1-\lambda)^{t} D_{0}$$

$$T_{t} = \lambda \sum_{i=1}^{t} (1-\lambda)^{t-i} D_{i} + (1-\lambda)^{t} T_{0}.$$
(11)

Substituting E_t into D_t , and thereafter substituting D_t into T_t , then after some algebraic manipulations, the charting statistic of the TEWMA *W* chart can also be re-written as

$$T_t = \frac{\lambda^3}{2} \sum_{j=1}^t (1-\lambda)^{t-j} (t-j+1)(t-j+2)W_j + \left(\frac{(1-\lambda)^t}{2}\right) (\lambda t(\lambda t + \lambda + 2) + 2)T_0.$$
(12)

For a more detailed algebraic derivation in the case of the parametric TEWMA \overline{X} chart, readers are advised to consult Alevizakos et al³⁶. Hence, the mean and variance of (9) or (12) are respectively given by

$$E(T_t) = \mu_T = \mu_W$$

(13)

and

$$Var(T_t) = \sigma_{T_t}^2 = \left(\frac{\lambda^6}{4} \sum_{j=1}^t (t-j+1)^2 (t-j+2)^2 (1-\lambda)^{2(t-j)}\right) \sigma_W^2.$$

The simplified version of the $Var(T_t)$ is given in the Appendix. The upper and lower control limits are respectively given by

$$UCL_t/LCL_t = \mu_T \pm L_T \sigma_{T_t},\tag{14}$$

where $L_T > 0$ is the control limits coefficient which has to be calculated such that the attained ARL_0 is close or equal to the pre-specified nominal ARL_0 . Thus, the TEWMA *W* chart gives an OOC signal when T_t value plots beyond the control limits in (14).

3.2 TEWMA W chart with a basic FIR feature

To improve the sensitivity of the standard TEWMA W chart in detecting shifts during startup, the basic FIR feature is added. This feature can be implemented to get an early alert at the beginning period of monitoring (when a change is suspected to be present) or after an OOC signal. The basic FIR feature proposed by Steiner⁴² is computed as

$$FIR(t) = 1 - (1 - f)^{(1+a)(t-1)},$$
(15a)

where it is suggested to use f = 0.5 and a = 0.3, with f as a segment of an interval from a start-up value to the control limits. Based on the latter formulation, Steiner⁴² showed that the FIR feature has 'very little to none' effect in narrowing control limit of the time-varying

control chart after the 20^{th} observation. Combining (14) and (15a) yields the control limits of the FIR-TEWMA *W* chart:

$$UCL_t/LCL_t = \mu_T \pm L_T \times FIR(t) \times \sigma_{T_t}.$$
 (15b)

At time t, the FIR-TEWMA W chart produces an OOC signal when T_t value plots beyond the control limits in (15b).

3.3 TEWMA W chart with a MFIR feature

As an improvement to Steiner⁴²'s FIR feature, Haq et al⁴³ introduced the MFIR feature. The MFIR is computed as

MFIR(t) =
$$(1 - (1 - f)^{(1+a)(t-1)})^{(1+\frac{1}{t})}$$
. (16a)

Combining (14) and (16a) yields the control limits of the MFIR-TEWMA W chart:

$$UCL_t/LCL_t = \mu_T \pm L_T \times MFIR(t) \times \sigma_{T_t}.$$
 (16b)

The MFIR-TEWMA W chart produces an OOC signal at time t when T_t value plots beyond the control limits in (16b). Haq et al⁴³ similarly showed that after the 20th observation, the MFIR feature has a 'very little to none' effect on narrowing the control limits of the timevarying control chart.

3.4 TEWMA W chart with a IMFIR feature

This paper also proposes a new FIR feature called the IMFIR feature where the power transformation factor is multiplied by square root of time. As it will be seen, the advantage of this new feature is that the sensitivity of the proposed TEWMA W chart increases the ability to detect any shifts at start-up period of the process as soon as possible. That is, it allows time-varying control limits to be narrower at the early stages of the process to better detect start-up problems, while ensuring that the long-run properties of the time-varying monitoring scheme are not disturbed. The IMFIR feature is defined by

IMFIR
$$(t) = \left(1 - (1 - f)^{(1+a)(t-1)}\right)^{\sqrt{t}\left(1 + \frac{1}{t}\right)},$$
 (17a)

where the values of f and a are taken as 0.5 and 0.3, respectively. Combining (14) and (17a) yields the control limits of the IMFIR-TEWMA W chart:

$$UCL_t/LCL_t = \mu_T \pm L_T \times \text{IMFIR}(t) \times \sigma_{T_t}.$$
 (17b)

The IMFIR-TEWMA W chart produces an OOC signal when T_t value plots beyond the control limits in (17b).

In Figure 1, the adjustment factors in (15a), (16a) and (17a) are compared against each other as time increases. This is an extension of Figure 1 in Haq et al⁴³ to illustrate how the new IMFIR feature is an improvement of both FIR and MFIR features by Steiner⁴² and Haq et

al⁴³. The adjustment factors of FIR, MFIR and IMFIR converge to unity as time increases; based on Figure 1, it is apparent that those of FIR and MFIR does so at time point equal to 24, while that of IMFIR does so at time point equal to 32. Moreover, while the adjustment factor of MFIR is smaller than that of FIR whenever $t \le 15$ and thereafter, they are equal; that of IMFIR is only equal to those of FIR and MFIR at sampling time 32 and onwards. This implies that the IMFIR feature reduces the control limits much further than the FIR and MFIR features at the start-up period (while not disturbing its long-run properties), and thus increase the detection capability of a time-varying TEWMA scheme.



Figure 1: Comparison of the FIR, MFIR and IMFIR adjustment factors with respect to time

3.5 The implementation of IMFIR-TEWMA W charts

The IMFIR-TEWMA W chart can be implemented as follows:

- Step 1: Generate a phase I sample of size m, X, from an IC process (e.g., from N(0,1)).
- Step 2: Generate a phase II sample of size *n*, *Y*. When the process is IC, the distributions of the phases I and II are identical (i.e. $\delta = 0$). Otherwise, the distribution for the phase II sample is taken to be similar to that for the phase I sample, but with a change in the mean parameter ($\delta \neq 0$) (e.g. from $N(\delta, 1)$).
- Step 3: The W statistic is computed by combining the phase I and phase II samples using (1).

- Step 4: Find the mean and variance of W when the process is IC using (2). These are used in Steps 5 and 6.
- Step 5: Compute the plotting statistic using (9) or (12) and choose L_T such that the attained $ARL_0 \approx 500$. The time-varying control limits of the IMFIR-TEWMA chart are calculated using (17b).
- Step 6: The IMFIR-TEWMA W chart gives a signal if the plotting statistic computed is Step 5 plots beyond the control limits calculated in Step 5.

The FIR-TEWMA and MFIR-TEWMA charts can also be implemented in a similar way.

4. Performance analysis and comparisons of the proposed control charts

4.1 Performance metrics

The run-length (*RL*), *ARL*, standard deviation of the *RL* (*SDRL*) and expected *ARL* (*EARL*) characteristics are often used to explore how sensitive is the explored control chart. The *RL* refers to the number of charting statistics to be plotted on the chart before the first OOC signal. In this paper, Monte Carlo simulations are used with 20 000 replications to assess the TEWMA schemes with and without FIR features. When the overall performance is of interest, the *EARL* metric is recommended. The *EARL* is mathematically defined by

$$EARL_{(\delta_{min}, \delta_{max}]} = \frac{1}{\Delta} \sum_{\delta = \delta_{min}}^{\delta_{max}} ARL(\delta)$$
(18)

where Δ is the number of increments between the δ_{min} (lower limit of δ) and δ_{max} (upper limit of δ).

4.2 IC robustness

To study the robustness of the proposed charts, three distributions are considered, i.e., the standard normal (denoted as N(0,1)), Student's t(v) (with v = 5, 15 and 30), Gamma (β ,1) (denoted by $GAM(\beta,1)$ with $\beta = 1$, 3 and 15) distributions are used. A control chart is considered to be IC robust when the characteristics of the IC *RL* are approximately the same across all continuous probability distributions (see for instance Chakraborti and Graham¹). Table 1 shows the robustness of the TEWMA *W* charts with and without FIR features when (m, n) = (100, 5) and $\lambda \in \{0.05, 0.5\}$ for a nominal $ARL_0=500$. From Table 1, it can be noticed that, the attained IC ARL_0 values are close to each other across different distributions and similarly, the attained IC $SDRL_0$ are also close to each other across different distributions considered in this paper. In addition, it is observed that the TEWMA *W* chart with no FIR

feature has lower $SDRL_0$ values than the ones with FIR features. Note that 5th, 25th, 50th, 75th and 95th percentiles are denoted as P5, P25, P50, P75 and P95, respectively and are presented in Figure 1. Moreover, it is shown in Figure 2 that the ARL_0 , P5, P25, P50 and P75 of the *RL* distribution are almost constant, while the ones of the $SDRL_0$ and P95 have a slight (but insignificant) variation. Thus, based on Table 1 and Figure 2, it can be deduced that the TEWMA *W* chart with or without FIR features is IC robust. It is important to mention that the *SDRL* and P95 profiles of the TEWMA *W* with the IMFIR feature are larger than the one with no FIR feature, see Figures 2(a) and (b); a high IC *SDRL* is undesired as it may lead to a high probability of an OOC signal, when in fact the process is IC.

Table 1. Robustness of the TEWMA *W* with and without FIR feature when (m, n) = (100, 5), $\lambda \in \{0.05, 0.5\}$ for a nominal $ARL_0 = 500$ under different distributions

		TE	WMA	FIR-T	EWMA	MFIR-7	ГEWMA	IMFIR-	TEWMA	
λ	Distribution	ARL ₀	SDRL ₀							
	N(0,1)	500.3	961.3	502.9	1166.7	500.8	1697.0	499.3	1506.6	
	<i>t</i> (5)	496.2	939.3	502.0	1128.6	514.0	1611.0	497.3	1580.4	
	<i>t</i> (15)	508.2	961.6	506.5	1163.4	501.3	1670.9	495.5	1480.5	
0.05	<i>t</i> (30)	496.5	941.2	507.7	1173.2	519.3	1660.0	504.0	1807.9	
	GAM(1,1)	501.4	956.3	503.8	1138.0	489.5	1616.6	499.1	1638.8	
	<i>GAM</i> (3,1)	500.4	945.8	500.4	1142.6	517.0	1657.3	505.4	1710.0	
	<i>GAM</i> (15,1)	510.1	950.8	502.8	1125.5	502.1	1563.4	500.7	1619.7	
	L_T	2.	321	2.	2.424		624	2.	617	
	<i>N</i> (0,1)	500.2	725.0	498.5	870.7	506.2	1228.2	494.4	1592.5	
	<i>t</i> (5)	507.1	735.6	508.2	879.6	498.7	1152.5	511.0	1631.2	
	<i>t</i> (15)	507.1	778.6	496.0	867.5	503.0	1240.2	511.2	1635.5	
0.5	<i>t</i> (30)	501.6	739.4	508.2	882.4	500.5	1171.9	505.5	1591.6	
	GAM(1,1)	495.7	702.9	505.9	884.7	507.6	1192.5	509.4	1626.7	
	<i>GAM</i> (3,1)	495.5	705.6	496.4	896.6	502.0	1181.0	508.7	1623.8	
	GAM(15,1)	511.6	729.0	501.0	871.2	508.9	1229.2	498.5	1529.5	
	L_T	2.	933	2.	2.995		126	3.210		



(a) TEWMA *W* chart (with no FIR feature) (b) IMFIR-TEWMA *W* chart **Figure 2**. Robustness of the proposed charts when (m, n) = (100,5), $\lambda = 0.05$ for a nominal $ARL_0=500$ under different distributions

4.3 OOC performance

Table 2 shows that the OOC ARL of the TEWMA W chart are lower for small values of λ , but increase as λ increases, indicating a deteriorating performance. For small and moderate shifts in the process location, the proposed chart performs better under the heavy-tailed tdistribution followed by the normal distribution. However, for large shifts, their performances under the different distributions are almost similar. For skewed gamma distribution, the higher the value of λ , the worse the performance; for instance, when $\delta = 0.1$, $\lambda = 0.9$, under the GAM(3,1) distribution, the OOC ARL is equal to 534.8 (i.e. higher than the nominal ARL₀ value of 500, which implies that it is ARL-biased). Next, the EARL values increase as λ increases for each distribution in Table 2, indicating a deteriorating OOC performance. For instance, under the N(0,1) distribution, when $\lambda = 0.05$, 0.5 and 0.9, then the corresponding EARL values are equal to 38.3, 45.4 and 61.7, respectively. In addition, Figure 3 studies the effect of Phase II sample size, i.e. $n \in \{3,5,10\}$, on the OOC performance using the EARLs when $\lambda \in \{0.05, 0.5\}$. Figure 3 shows that as the Phase II sample size increases, the better the Phase II OOC performance becomes, thus indicating an improvement in the detection ability. Similarly, Figure 4 shows that the larger the Phase I sample size (m), the better the Phase II OOC performance. Figures 3 and 4 show that the larger the smoothing parameter, the less sensitive is the TEWMA W chart, since the EARLs become larger; especially for the skewed gamma distribution.

	distributions										
		<i>N</i> (0,1)			<i>t</i> (5)		<i>GAM</i> (3,1)				
Shift	$\lambda = 0.05$	$\lambda = 0.5$	$\lambda = 0.9$	$\lambda = 0.05$	$\lambda = 0.5$	$\lambda = 0.9$	$\lambda = 0.05$	$\lambda = 0.5$	$\lambda = 0.9$		
0.1	355.3	369.1	408.3	304.5	360.0	401.9	339.9	435.2	534.8		
0.2	126.5	176.5	244.4	93.3	137.6	224.7	110.1	192.1	357.9		
0.3	36.7	63.6	121.9	24.1	39.9	103.9	28.9	63.9	184.6		
0.4	16.4	24.6	61.0	11.8	16.3	46.4	13.2	21.5	82.2		
0.5	9.7	12.2	32.8	7.5	8.9	23.4	7.6	10.3	42.1		
0.6	6.7	7.8	17.7	5.3	5.9	12.7	5.3	6.3	23.1		
0.7	4.9	5.4	10.9	3.9	4.5	7.8	3.9	4.6	12.4		
0.8	3.8	4.3	7.3	3.1	3.5	5.3	3.0	3.6	7.9		
0.9	3.1	3.5	5.3	2.5	2.9	3.8	2.4	3.0	5.4		
1.0	2.5	3.0	4.0	2.1	2.5	3.0	2.0	2.6	3.9		
1.1	2.1	2.6	3.2	1.8	2.3	2.5	1.7	2.3	2.9		
1.2	1.9	2.3	2.6	1.6	2.0	2.2	1.5	2.1	2.4		
1.3	1.6	2.1	2.3	1.4	1.8	1.8	1.4	1.9	2.1		
1.4	1.5	1.9	2.0	1.3	1.7	1.7	1.2	1.8	1.8		
1.5	1.3	1.7	1.8	1.2	1.5	1.5	1.1	1.6	1.6		
EARL(0,1.5]	38.3	45.4	61.7	31.0	39.4	56.2	34.9	50.2	84.3		
L_T	2.321	2.933	2.851	2.321	2.933	2.851	2.321	2.933	2.851		

Table 2. The *ARL* profile of the TEWMA *W* chart when $(m, n) = (100, 5), \lambda \in \{0.05, 0.5, 0.9\}, \delta_{\min}=0$ and $\delta_{\max}=1.5$ with an increment of 0.1 for a nominal *ARL*₀=500 under different distributions



Figure 3. The effect of Phase II sample size on the performance TEWMA *W* chart using the *EARL* when $m = 100, n \in \{3, 5, 10\}, \lambda \in \{0.05, 0.5\}, \delta_{\min}=0$ and $\delta_{\max}=2$ with an increment of 0.1



Figure 4. The effect of Phase I sample size on the performance TEWMA *W* chart using the *EARL* when $m = 100, m \in \{50, 100, 500\}, \lambda \in \{0.05, 0.5\}, \delta_{\min}=0$ and $\delta_{\max}=2$ with an increment of 0.1

At each shift value and for different *EARLs*, the best performing chart is boldfaced in Tables 4 to 6. Note that $EARL_{(0,0.5]}$, $EARL_{(1,1.5]}$ and $EARL_{(0,1.5]}$ denote overall performance metrics defined in (18) for 'small', 'large' and 'small-to-large' shifts, respectively. Based on the OOC *ARL* performance illustrated in Table 3, it shows that the IMFIR-TEWMA *W* chart is more sensitive than the standard TEWMA, FIR-TEWMA and MFIR-TEWMA *W* charts for all shift values as well as in terms of the *EARL* values under symmetric distributions. However, for small shifts under skewed distributions, the IMFIR-TEWMA *W* chart is outperformed by its competitors, see Table 3 when $\delta = 0.1$ under the *GAM*(3,1) distribution. For symmetric distributions, the IMFIR-TEWMA *W* chart has a better overall performance than the other charts, see the *EARL* values under the *N*(0,1) and *t*(5) distributions. However, for the skewed gamma distribution, the FIR-TEWMA *W* chart is slightly better than the other charts. It is important to note that, in Table 3, the TEWMA *W* chart with no FIR feature has the worst performance than those with FIR, MFIR or IMFIR.

		N	0,1)			t	(5)		<i>GAM</i> (3,1)				
Shift	TEWMA	FIR- TEWMA	MFIR- TEWMA	IMFIR- TEWMA	TEWMA	FIR- TEWMA	MFIR- TEWMA	IMFIR- TEWMA	TEWMA	FIR- TEWMA	MFIR- TEWMA	IMFIR- TEWMA	
0.1	369.1	371.8	373.9	357.5	360.0	351.5	340.6	328.2	435.2	428.9	452.9	478.5	
0.2	176.5	170.1	157.8	142.0	137.7	127.2	119.2	98.6	192.1	203.5	197.8	187.1	
0.3	63.6	51.7	47.2	35.7	40.0	30.5	28.2	18.0	63.9	59.0	56.6	39.4	
0.4	24.6	15.6	13.8	9.5	16.3	9.7	7.2	4.6	21.5	14.2	11.9	14.2	
0.5	12.2	6.6	4.9	3.0	9.0	4.7	3.0	2.1	10.3	5.6	3.8	2.3	
0.6	7.8	4.0	2.7	1.9	5.9	2.8	1.9	1.6	6.3	3.0	2.1	1.6	
0.7	5.4	2.7	1.9	1.5	4.5	2.1	1.5	1.3	4.6	2.1	1.5	1.3	
0.8	4.3	2.0	1.5	1.3	3.5	1.7	1.3	1.2	3.6	1.6	1.2	1.1	
0.9	3.5	1.7	1.3	1.2	3.0	1.4	1.2	1.1	3.0	1.3	1.1	1.1	
1.0	3.0	1.4	1.2	1.1	2.5	1.3	1.1	1.1	2.6	1.2	1.0	1.0	
1.1	2.6	1.3	1.1	1.1	2.3	1.2	1.1	1.1	2.3	1.1	1.0	1.0	
1.2	2.3	1.2	1.1	1.0	2.0	1.1	1.0	1.0	2.1	1.1	1.0	1.0	
1.3	2.1	1.1	1.0	1.0	1.8	1.1	1.0	1.0	1.9	1.0	1.0	1.0	
1.4	1.9	1.1	1.0	1.0	1.7	1.1	1.0	1.0	1.8	1.0	1.0	1.0	
1.5	1.7	1.1	1.0	1.0	1.6	1.0	1.0	1.0	1.6	1.0	1.0	1.0	
$EARL_{(0,0.5]}$	129.2	123.2	119.5	109.5	112.6	104.7	99.6	90.3	144.6	142.2	144.6	144.3	
EARL _{(1,1.5]}	2.1	1.2	1.0	1.0	1.9	1.1	1.0	1.0	1.9	1.0	1.0	1.0	
EARL(0,1.5]	45.4	42.2	40.8	37.3	39.5	35.9	34.0	30.9	50.2	48.4	49.0	48.8	
L_T	2.933	2.995	3.126	3.2095	2.933	2.995	3.126	3.2095	2.933	2.995	3.126	3.2095	

Table 3. The *ARL* profile of the TEWMA *W* with and without FIR features when $(m, n, \lambda) = (100, 5, 0.5)$ for a nominal *ARL*₀=500 under different distributions

Since the *ARL* values are spread over a wide range of values; hence, in Figure 5, they are displayed in a compact manner using the log scale, i.e. $\log_{10}(ARL)$. From Figure 5, it is observed that, in general, the standard TEWMA *W* chart will take longer to give a signal as compared to the FIR-TEWMA, MFIR-TEWMA and IMFIR-TEWMA *W* charts for the majority of shifts. However, for most shifts, the IMFIR-TEWMA *W* chart detects changes in the location parameter faster, with the MFIR-TEWMA and FIR-TEWMA *W* charts, being the second and third best for the majority of shifts, respectively.



Figure 5. The OOC $\log_{10}(ARL)$ profile comparison of the proposed schemes when $(m, n) = (100, 5), \lambda = 0.05$ for a nominal $ARL_0=500$ under different distributions

4.4 Performance comparison of the EWMA, DEWMA and TEWMA W charts

In Tables 4 to 6, the proposed TEWMA *W* chart is compared to the EWMA and DEWMA *W* charts with and without FIR features when $\lambda = 0.5$, m = 100 and n = 5 under the N(0,1), t(5)

and GAM(3,1), respectively. It is worth mentioning that the EWMA and DEWMA W charts discussed in Li et al²⁶ and Malela-Majika²⁸ did not incorporate any of the FIR features discussed here. For control charts with no FIRs (i.e. labelled 'Standard' in Tables 4 to 6), the TEWMA W chart has the best small as well as large shifts performance; however, the DEWMA W chart has better moderate shifts performance. With respect to *EARLs*, the TEWMA W chart has the best performance followed by the DEWMA W chart and finally, the EWMA W chart being the worst performing.

Next, for control charts with FIR and MFIR features, the DEWMA and TEWMA W charts seem to have the best small shifts performance, with the TEWMA chart having a slight advantage based on the $EARL_{(0,0.5]}$ values. Note though, for moderate shifts, the DEWMA W chart has better performance. For large shifts, i.e. for $\delta \ge 1$, the competing charts are almost equivalent, see the $EARL_{(1,1.5]}$ values in Tables 4 to 6.

Finally, for control charts with IMFIR feature, the TEWMA *W* chart seems to have the best small shifts performance, see the $EARL_{(0,0.5]}$ values in Tables 4 to 6; with the DEWMA *W* chart having the second best performance. Again, for moderate shifts, the DEWMA *W* chart has a better performance. For large shifts, all the competing charts are almost equivalent; see the $EARL_{(1,1.5]}$ values in Tables 4 to 6.

In general, from Tables 4 to 6, it is observed that in most situations (with or without the FIR features), the TEWMA *W* chart has relatively better OOC performance for small and large shifts; however, the DEWMA *W* chart has excellent moderate shifts detection ability than its competitors. Moreover, comparing the EWMA, DEWMA and TEWMA *W* charts' *ARL* and *EARL* values from the columns, 'Standard' vs. 'FIR' vs. 'MFIR' vs. 'IMFIR': the feature that results in the lowest *ARL* and *EARL* values, correspond to the IMFIR. The latter shows that for each chart, the IMFIR feature yields the best performance in most situations.

	Standard			FIR				MFIR		IMFIR		
Shift	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA
0.1	357.7	354.3	352.7	353.2	335.8	349.1	340.4	327.5	340.8	340.2	326.1	333.6
0.2	142.5	128.5	118.0	132.3	117.6	120.6	116.6	116.5	108.2	117.7	110.9	99.2
0.3	38.9	36.6	37.5	32.9	33.1	32.6	26.4	25.9	25.1	27.0	26.6	23.7
0.4	15.6	14.2	15.6	12.7	11.7	13.0	9.2	9.0	9.1	9.1	8.1	9.3
0.5	9.3	8.8	9.8	7.3	7.0	7.8	5.1	4.2	5.4	5.2	5.1	5.4
0.6	6.5	6.1	6.7	5.0	4.7	5.2	3.3	2.8	3.8	3.4	3.3	3.6
0.7	5.0	4.7	5.1	3.5	3.4	3.7	2.4	2.0	2.7	2.5	2.4	2.6
0.8	4.0	3.6	3.8	2.7	2.6	2.7	1.8	1.5	2.0	1.9	1.9	2.0
0.9	3.3	3.0	3.0	2.2	2.0	2.1	1.5	1.3	1.6	1.5	1.5	1.5
1.0	2.9	2.5	2.5	1.8	1.7	1.7	1.3	1.2	1.4	1.3	1.3	1.3
1.1	2.5	2.1	2.1	1.5	1.5	1.4	1.2	1.2	1.2	1.2	1.2	1.2
1.2	2.2	1.9	1.9	1.3	1.3	1.2	1.1	1.1	1.1	1.1	1.1	1.1
1.3	2.0	1.7	1.6	1.2	1.2	1.2	1.1	1.0	1.1	1.1	1.1	1.1
1.4	1.9	1.5	1.5	1.1	1.1	1.1	1.0	1.0	1.0	1.0	1.0	1.0
1.5	1.7	1.4	1.3	1.1	1.1	1.1	1.0	1.0	1.0	1.0	1.0	1.0
EARL(0,0.5]	112.8	108.5	106.7	107.7	101.0	104.6	99.5	96.6	97.7	99.8	95.4	94.2
EARL _{(1,1.5]}	2.1	1.7	1.7	1.3	1.2	1.2	1.1	1.1	1.1	1.1	1.1	1.1
EARL(0,1.5]	39.7	38.1	37.5	37.3	35.0	36.3	34.2	33.1	33.7	34.3	32.8	32.5
	2.906	2.474	2.318	2.991	2.575	2.423	3.135	2758	2.624	3.12	2.75	2.623

Table 4. The OOC *ARL* profile comparison of the EWMA, DEWMA and TEWMA charts with and without FIR features when $(m, n, \lambda) = (100, 5, 0.5)$ for a nominal *ARL*₀=500 under the *N*(0,1) distribution

Table 5. The OOC *ARL* profile comparison of the EWMA, DEWMA and TEWMA charts with and without FIR features when $(m, n, \lambda) = (100, 5, 0.5)$ for a nominal *ARL*₀=500 under the *t*(5) distribution

	Standard				FIR		MFIR			IMFIR			
Shift	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	
0.1	337.5	318.7	306.3	322.3	316.1	308.7	320.1	300.1	315.3	324.5	293.8	277.4	
0.2	97.1	87.5	88.0	93.0	77.9	77.8	88.8	75.5	75.3	86.1	74.7	81.8	
0.3	26.5	22.2	25.1	21.4	21.5	23.3	15.2	16.4	13.9	17.2	15.6	15.9	
0.4	11.3	10.6	12.1	9.6	8.7	10.4	6.4	6.1	6.7	6.2	6.1	6.7	
0.5	7.3	6.7	7.5	5.8	5.4	6.0	3.8	3.7	4.2	3.9	3.7	4.1	
0.6	5.3	4.8	5.3	3.8	3.7	4.0	2.6	2.6	2.9	2.6	2.5	2.8	
0.7	4.1	3.7	4.0	2.8	2.7	2.8	2.0	1.9	2.1	1.9	1.9	2.0	
0.8	3.3	2.9	3.1	2.2	2.1	2.1	1.5	1.6	1.6	1.5	1.5	1.5	
0.9	2.8	2.4	2.5	1.8	1.7	1.7	1.3	1.3	1.4	1.3	1.3	1.3	
1.0	2.5	2.1	2.1	1.5	1.4	1.4	1.2	1.2	1.2	1.2	1.2	1.2	
1.1	2.2	1.8	1.8	1.3	1.3	1.3	1.1	1.1	1.1	1.1	1.1	1.1	
1.2	2.0	1.6	1.6	1.2	1.2	1.2	1.1	1.1	1.1	1.1	1.1	1.1	
1.3	1.8	1.5	1.4	1.1	1.1	1.1	1.0	1.0	1.0	1.0	1.0	1.0	
1.4	1.6	1.4	1.3	1.1	1.1	1.1	1.0	1.0	1.0	1.0	1.0	1.0	
1.5	1.5	1.3	1.3	1.1	1.1	1.0	1.0	1.0	1.0	1.0	1.0	1.0	
$EARL_{(0,0.5]}$	95.9	89.2	87.8	90.4	85.9	85.2	86.9	80.4	83.1	87.6	78.8	77.2	
EARL(1,1.5)	1.8	1.5	1.5	1.2	1.1	1.1	1.1	1.1	1.1	1.0	1.0	1.0	
EARL(0,1.5]	33.8	31.3	30.9	31.3	29.8	29.6	29.9	27.7	28.7	30.1	27.2	26.7	
L_T	2.906	2.474	2.318	2.991	2.575	2.423	3.135	2758	2.624	3.12	2.75	2.623	

_		Standard			FIR				MFIR		IMFIR			
	Shift	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	EWMA	DEWMA	TEWMA	
	0.1	356.8	345.5	332.4	363.4	356.2	343.9	375.3	370.5	354.0	362.2	355.5	343.6	
	0.2	133.3	111.3	110.0	131.8	112.7	110.0	109.5	101.3	104.8	120.8	116.6	116.5	
	0.3	31.7	30.4	27.7	35.2	24.6	24.4	25.4	20.9	21.4	31.2	24.5	21.0	
	0.4	12.2	11.9	12.2	10.8	9.4	10.1	8.5	6.9	7.8	8.4	6.8	7.1	
	0.5	7.5	7.0	7.6	6.1	5.5	6.0	4.0	3.8	4.2	3.9	3.8	4.2	
	0.6	5.4	4.8	5.3	4.0	3.6	3.8	2.5	2.4	2.7	2.5	2.5	2.6	
	0.7	4.2	3.7	3.9	2.9	2.6	2.6	1.8	1.7	1.8	1.8	1.7	1.8	
	0.8	3.4	2.9	3.0	2.1	1.9	1.9	1.4	1.3	1.4	1.4	1.3	1.3	
	0.9	2.9	2.4	2.4	1.6	1.5	1.5	1.2	1.2	1.2	1.2	1.2	1.1	
	1.0	2.5	2.1	2.0	1.4	1.3	1.2	1.1	1.1	1.1	1.1	1.1	1.1	
	1.1	2.2	1.8	1.7	1.2	1.1	1.1	1.0	1.0	1.0	1.0	1.0	1.0	
	1.2	2.0	1.6	1.5	1.1	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	
	1.3	1.8	1.4	1.3	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	
	1.4	1.7	1.3	1.2	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	
	1.5	1.6	1.2	1.1	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	
	EARL(0,0.5]	108.3	101.2	98.0	109.4	101.7	98.9	104.5	100.7	98.4	105.3	101.4	98.5	
	EARL(1,1.5]	1.9	1.5	1.4	1.1	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	
	EARL(0,1.5]	37.9	35.3	34.2	37.6	35.0	34.0	35.7	34.4	33.7	36.0	34.7	33.7	
	L_T	2.906	2.474	2.318	2.991	2.575	2.423	3.135	2758	2.624	3.12	2.75	2.623	

Table 6. The OOC *ARL* profile comparison of the EWMA, DEWMA and TEWMA charts with and without FIR features when $(m, n, \lambda) = (100, 5, 0.5)$ for a nominal *ARL*₀=500 under the *GAM*(3,1) distribution

5. Illustrative example

The implementation of the proposed TEWMA W control chart with and without the IMFIR feature using a dataset on the size of piston rings from Montgomery⁴⁵ for which a goodness of fit test for normality is not rejected since the p-value is equal to 0.4167. Setting an $ARL_0=500$, we use $\lambda=0.50$ and $L_T=2.937$, so that the attained ARL_0 value of the TEWMA W chart be equal to 500.14 for (m,n)=(125,5). The TEWMA W chart is shown in Figure 6 and no charting statistic plots beyond the control limits, i.e. it does not give a signal in the prospective phase. When the IMFIR feature is incorporated in the TEWMA W chart, it is found that $L_T = 3.177$ yields an attained ARL_0 value of 501.22. The IMFIR-TEWMA W chart is also shown in Figure 6 where we observe that it gives a signal on the first sample in the prospective phase. Moreover, it is observed in Figure 6 that, the IMFIR adjustment factor narrows the control limits during the start-up period. This in turn, ensures that the IMFIR-TEWMA W chart has a better detection ability than the TEWMA W chart (without the IMFIR feature). Also, it can be seen that, as the time increases, the control limits of the IMFIR-TEWMA W chart converges towards those of the TEWMA W chart without the IMFIR, which shows that the effect of the IMFIR feature is only applicable during start-up only and gradually disappears as time increases.



Figure 6. Illustrative example of the standard TEWMA and IMFIR-TEWMA *W* control charts using the piston ring data from Montgomery⁴⁵

6. Conclusion

This paper proposed the TEWMA chart based on the Wilcoxon rank-sum (W) statistic with and without the IMFIR feature. A comparison against other FIR features indicates that the IMFIR feature has the best detection ability for the new TEWMA W chart in most situations. More importantly, the TEWMA chart outperforms the DEWMA and EWMA W charts with and without the FIR features, especially for small shifts. However, the DEWMA W chart has the best performance in detecting moderate shifts. Therefore, it can be said that the proposed TEWMA W chart with IMFIR feature can be used over the DEWMA and EWMA W charts as it is shown that it is more sensitive and results in earlier OOC signals in most situations as compared to its competitors.

For future research, we intend to propose the TEWMA chart to monitor variability using the Ansari-Bradley statistic (refer for instance to Zhou et al⁴⁶) and also study the effect of the IMFIR feature.

Appendix

With the help of a formal calculus software, $Var(T_t)$ can equivalently be written as

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$$\begin{aligned} Var(T_t) &= \left(\frac{\lambda^6}{4} \sum_{j=1}^t (t-j+1)^2 (t-j+2)^2 (1-\lambda)^{2(t-j)}\right) \sigma_W^2 \\ &= \frac{1}{(\lambda-2)^2 \lambda^5} \left(\lambda^8 t^2 (t+1)^2 (1-\lambda)^{2t} - 8\lambda^7 t^2 (t+1)^2 (1-\lambda)^{2t} \\ &+ 4\lambda^6 t^2 (6t^2+11t+6) (1-\lambda)^{2t} - 8\lambda^5 t^2 (4t^2+5t+4) (1-\lambda)^{2t} \\ &+ 8\lambda^2 (6t^2 (1-\lambda)^{2t} - 9t (1-\lambda)^{2t} + 5(1-\lambda)^{2t} - 5) \\ &+ 16\lambda^3 (2t^3 (1-\lambda)^{2t} - 3t^2 (1-\lambda)^{2t} + 3t (1-\lambda)^{2t} - (1-\lambda)^{2t} + 1) \\ &+ 4\lambda^4 (4t^4 (1-\lambda)^{2t} - 4t^3 (1-\lambda)^{2t} + 7t^2 (1-\lambda)^{2t} - 3t (1-\lambda)^{2t} \\ &+ (1-\lambda)^{2t} - 1) + 24((1-\lambda)^{2t} - 1) \\ &+ 48\lambda (t(1-\lambda)^{2t} - (1-\lambda)^{2t} + 1) \right) \sigma_W^2. \end{aligned}$$

Another alternative formulation of the above expression is provided in Alevizakos et al³⁶.

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