

Nonparametric Estimation for I.I.D. Paths of a Martingale Driven Model with Application to Non-Autonomous Fractional SDE

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NONPARAMETRIC ESTIMATION FOR I.I.D. PATHS OF A MARTINGALE DRIVEN MODEL WITH APPLICATION TO NON-AUTONOMOUS FRACTIONAL SDE

NICOLAS MARIE†

ABSTRACT. This paper deals with a projection least square estimator of the function J_0 computed from multiple independent observations on [0,T] of the process Z defined by $dZ_t = J_0(t)d\langle M \rangle_t + dM_t$, where M is a centered, continuous and square integrable martingale vanishing at 0. Risk bounds are established on this estimator and on an associated adaptive estimator. An appropriate transformation allows to rewrite the differential equation $dX_t = V(X_t)(b_0(t)dt + \sigma(t)dB_t)$, where B is a fractional Brownian motion of Hurst parameter $H \in (1/2, 1)$, as a model of the previous type. So, the second part of the paper deals with risk bounds on a nonparametric estimator of b_0 derived from the results on the projection least square estimator of J_0 . In particular, our results apply to the estimation of the drift function in a non-autonomous extension of the fractional Black-Scholes model introduced in Hu et al. [15].

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1. Introduction

Since the 1980's, the statistical inference for stochastic differential equations (SDE) driven by a Brownian motion has been widely investigated by many authors in the parametric and in the nonparametric frameworks. Classically (see Kutoyants [18]), the estimators of the drift function are computed from one path of the solution to the SDE and converge when the time horizon T>0 goes to infinity. The existence and the uniqueness of the stationary solution to the SDE are then required, and obtained thanks to

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restrictive conditions on the drift function.

Since few years, a new type of parametric and nonparametric estimators is investigated; those computed from multiple independent observations on [0,T] of the SDE solution. Indeed, this functional data analysis problem is already studied in the parametric framework (see Ditlevsen and De Gaetano [14], Overgaard et al. [21], Picchini, De Gaetano and Ditlevsen [22], Picchini and Ditlevsen [23], Comte, Genon-Catalot and Samson [6], Delattre and Lavielle [10], Delattre, Genon-Catalot and Samson [9], Dion and Genon-Catalot [13], Delattre, Genon-Catalot and Larédo [8], etc.) and more recently in the nonparametric framework (see Comte and Genon-Catalot [4, 5], Della Maestra and Hoffmann [11], and Marie and Rosier [19]). In [4, 5], the authors extend to the diffusion processes framework the projection least squares estimators already well studied in the regression framework (see Cohen et al. [2] and Comte and Genon-Catalot [3]). Our paper deals with a nonparametric estimation problem close to this last one.

Consider the stochastic process $Z = (Z_t)_{t \in [0,T]}$, defined by

(1)
$$Z_t = \int_0^t J_0(s)d\langle M \rangle_s + M_t \; ; \; \forall t \in [0, T],$$

where $M = (M_t)_{t \in [0,T]} \neq 0$ is a centered, continuous and square integrable martingale vanishing at 0, and J_0 is an unknown function which belongs to $\mathbb{L}^2([0,T],d\langle M\rangle_t)$. By assuming that the quadratic variation $\langle M\rangle_t$ of M is deterministic for every $t \in [0,T]$, our paper deals with the estimator $\widehat{J}_{m,N}$ of J_0 minimizing the objective function

$$J \longmapsto \gamma_{m,N}(J) := \frac{1}{N} \sum_{i=1}^{N} \left(\int_{0}^{T} J(s)^{2} d\langle M^{i} \rangle_{s} - 2 \int_{0}^{T} J(s) dZ_{s}^{i} \right)$$

on a m-dimensional function space \mathcal{S}_m , where M^1, \ldots, M^N (resp. Z^1, \ldots, Z^N) are $N \in \mathbb{N}^*$ independent copies of M (resp. Z) and $m \in \{1, \ldots, N\}$. Precisely, risk bounds are established on $\widehat{J}_{m,N}$ and on the adaptive estimator $\widehat{J}_{\widehat{m},N}$, where

$$\widehat{m} = \arg\min_{m \in \mathcal{M}_N} \{ \gamma_{m,N}(\widehat{J}_{m,N}) + \operatorname{pen}(m) \} \quad \text{with} \quad \operatorname{pen}(.) := \mathfrak{c}_{\operatorname{cal}} \frac{\cdot}{N} \quad \text{and} \quad \mathcal{M}_N \subset \{1,\ldots,N\}.$$

Now, consider the differential equation

(2)
$$X_t = x_0 + \int_0^t V(X_s)(b_0(s)ds + \sigma(s)dB_s) \; ; \; t \in [0, T],$$

where $x_0 \in \mathbb{R}^*$, B is a fractional Brownian motion of Hurst parameter $H \in (1/2, 1)$, the stochastic integral with respect to B is taken pathwise (in Young's sense), and $V : \mathbb{R} \to \mathbb{R}$, $\sigma : [0, T] \to \mathbb{R}^*$ and $b_0 : [0, T] \to \mathbb{R}$ are regular enough. An appropriate transformation (see Subsection 4.1) allows to rewrite Equation (2) as a model of type (1) driven by the Molchan martingale which quadratic variation is t^{2-2H} for every $t \in [0, T]$. Our paper also deals with a risk bound on an estimator of b_0/σ derived from $\widehat{J}_{m,N}$. Finally, let us consider a financial market model in which the prices of the risky asset are modeled by the following equation of type (2):

(3)
$$S_t = S_0 + \int_0^t S_u \left(\left(\overline{b}_0(u) - \frac{\sigma^2}{2} u^{2H-1} \right) du + \sigma dB_u \right) \; ; \; t \in [0, T]$$

where $S_0, \sigma \in (0, \infty)$ and $\bar{b}_0 \in C^0([0, T]; \mathbb{R})$. This is a non-autonomous extension of the fractional Black-Scholes model defined in Hu et al. [15]. An estimator of \bar{b}_0 is derived from $\widehat{J}_{m,N}$ at Subsection 4.3. Up to our knowledge, only Comte and Marie [7] deals with a nonparametric estimator of the drift function computed from multiple independent observations on [0, T] of the solution to a fractional SDE.

At Section 2, a detailed definition of the projection least square estimator of J_0 is provided. Section 3 deals with risk bounds on $\widehat{J}_{m,N}$ and on the adaptive estimator $\widehat{J}_{\widehat{m},N}$. At Section 4, the results of Section 3 on the estimator of J_0 are applied to the estimation of b_0 in Equation (2) and then of \overline{b}_0 in Equation (3). Finally, at Section 5, some numerical experiments on Model (1) are provided when M is the Molchan martingale.

2. A projection least square estimator of the map J_0

In the sequel, the quadratic variation $\langle M \rangle = (\langle M \rangle_t)_{t \in [0,T]}$ of M fulfills the following assumption.

Assumption 2.1. The (nonnegative, increasing and continuous) process $\langle M \rangle$ is a deterministic function. For some results, $\langle M \rangle$ fulfills the following stronger assumption.

Assumption 2.2. There exists $\mu \in C^0((0,T];\mathbb{R}_+)$ such that $1/\mu$ is continuous from [0,T] into \mathbb{R}_+ , and such that

$$\langle M \rangle_t = \int_0^t \mu(s) ds \; ; \; \forall t \in [0, T].$$

2.1. **The objective function.** In order to define a least square projection estimator of J_0 , let us consider $N \in \mathbb{N}^*$ independent copies M^1, \ldots, M^N (resp. Z^1, \ldots, Z^N) of M (resp. Z), and the objective function $\gamma_{m,N}$ defined by

$$\gamma_{m,N}(J) := \frac{1}{N} \sum_{i=1}^{N} \left(\int_{0}^{T} J(s)^{2} d\langle M^{i} \rangle_{s} - 2 \int_{0}^{T} J(s) dZ_{s}^{i} \right)$$

for every $J \in \mathcal{S}_m$, where $m \in \{1, \dots, N\}$, $\mathcal{S}_m := \operatorname{span}\{\varphi_1, \dots, \varphi_m\}$ and $\varphi_1, \dots, \varphi_N$ are continuous functions from [0, T] into \mathbb{R} such that $(\varphi_1, \dots, \varphi_N)$ is an orthonormal family in $\mathbb{L}^2([0, T], dt)$.

Remark. Note that since $t \in [0,T] \mapsto \langle M \rangle_t$ is nonnegative, increasing and continuous, and since the φ_j 's are continuous from [0,T] into \mathbb{R} , the objective function $\gamma_{m,N}$ is well-defined.

For any $J \in \mathcal{S}_m$,

$$\mathbb{E}(\gamma_{m,N}(J)) = \int_0^T J(s)^2 d\langle M \rangle_s - 2 \int_0^T J(s) J_0(s) d\langle M \rangle_s - 2 \mathbb{E}\left(\int_0^T J(s) dM_s\right)$$
$$= \int_0^T (J(s) - J_0(s))^2 d\langle M \rangle_s - \int_0^T J_0(s)^2 d\langle M \rangle_s.$$

Then, the more J is close to J_0 , the more $\mathbb{E}(\gamma_{m,N}(J))$ is small. For this reason, the estimator of J_0 minimizing $\gamma_{m,N}$ is studied in this paper.

2.2. The projection least square estimator. Consider the estimator

(4)
$$\widehat{J}_{m,N} := \arg \min_{J \in \mathcal{S}_m} \gamma_{m,N}(J)$$

of J_0 . Since $S_m = \text{span}\{\varphi_1, \dots, \varphi_m\}$, there exist m square integrable random variables $\widehat{\theta}_1, \dots, \widehat{\theta}_m$ such that

$$\widehat{J}_{m,N} = \sum_{j=1}^{m} \widehat{\theta}_j \varphi_j.$$

Then,

$$\nabla \gamma_{m,N}(\widehat{J}_{m,N}) = \left(\frac{1}{N} \sum_{i=1}^{N} \left(2 \sum_{k=1}^{m} \widehat{\theta}_{k} \int_{0}^{T} \varphi_{j}(s) \varphi_{k}(s) d\langle M^{i} \rangle_{s} - 2 \int_{0}^{T} \varphi_{j}(s) dZ_{s}^{i}\right)\right)_{j \in \{1,\dots,m\}}.$$

Therefore, by (4), necessarily

$$\widehat{\theta}_{m,N} := (\widehat{\theta}_1, \dots, \widehat{\theta}_m)^* = \mathbf{\Psi}_m^{-1} \mathbf{z}_{m,N},$$

where

$$\Psi_m := \left(\int_0^T \varphi_j(s) \varphi_k(s) d\langle M \rangle_s \right)_{j,k \in \{1,\dots,m\}} \quad \text{and} \quad \mathbf{z}_{m,N} := \left(\frac{1}{N} \sum_{i=1}^N \int_0^T \varphi_j(s) dZ_s^i \right)_{j \in \{1,\dots,m\}}.$$

3. Risk bound and model selection

In the sequel, the space $\mathbb{L}^2([0,T],d\langle M\rangle_t)$ is equipped with the scalar product $\langle .,.\rangle_{\langle M\rangle}$ defined by

$$\langle \varphi, \psi \rangle_{\langle M \rangle} := \int_0^T \varphi(s) \psi(s) d\langle M \rangle_s$$

for every $\varphi, \psi \in \mathbb{L}^2([0,T], d\langle M \rangle_t)$. The associated norm is denoted by $\|.\|_{\langle M \rangle}$.

First, the following proposition provides a risk bound on $\widehat{J}_{m,N}$ for a fixed $m \in \{1,\ldots,N\}$.

Proposition 3.1. Under Assumption 2.1,

(5)
$$\mathbb{E}(\|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2) \leqslant \min_{J \in \mathcal{S}_m} \|J - J_0\|_{\langle M \rangle}^2 + \frac{2m}{N}.$$

Note that Inequality (5) says first that the bound on the variance of our least square estimator of J_0 is of order m/N, as in the usual nonparametric regression framework. Under Assumption 2.2, the following corollary provides a more understandable expression of the bound on the bias in Inequality (5).

Corollary 3.2. Under Assumption 2.2,

$$\mathbb{E}(\|\widehat{J}_{m,N} - J_0\|_2^2) \leqslant \|\mu^{-1}\|_{\infty,T} \|p_{\mathcal{S}_m(\mu)}^{\perp}(\mu^{1/2}J_0) - \mu^{1/2}J_0\|_2^2 + 2\|\mu^{-1}\|_{\infty,T} \frac{m}{N}$$

where

$$S_m(\mu) := \{ \iota \in \mathbb{L}^2([0,T], dt) : \exists \varphi \in S_m, \forall t \in (0,T], \iota(t) = \mu(t)^{1/2} \varphi(t) \}.$$

For instance, assume that $S_m = \text{span}\{\overline{\varphi}_1, \dots, \overline{\varphi}_m\}$, where

$$\overline{\varphi}_1(t) := \sqrt{\frac{1}{\mu(t)T}}, \quad \overline{\varphi}_{2j}(t) := \sqrt{\frac{2}{\mu(t)T}} \cos\left(2\pi j \frac{t}{T}\right) \quad \text{and} \quad \overline{\varphi}_{2j+1}(t) := \sqrt{\frac{2}{\mu(t)T}} \sin\left(2\pi j \frac{t}{T}\right)$$

for every $t \in [0,T]$ and $j \in \mathbb{N}^*$ satisfying $2j+1 \leqslant m$. The basis $(\varphi_1,\ldots,\varphi_m)$ of \mathcal{S}_m , orthonormal in $\mathbb{L}^2([0,T],dt)$, is obtained from $(\overline{\varphi}_1,\ldots,\overline{\varphi}_m)$ via the Gram-Schmidt process. Consider also the Sobolev space

$$\mathbb{W}_{2}^{\beta}([0,T]) := \left\{ \iota \in C^{\beta-1}([0,T];\mathbb{R}) : \int_{0}^{T} \iota^{(\beta)}(t)^{2} dt < \infty \right\} \; ; \; \beta \in \mathbb{N}^{*},$$

and assume that there exists $\iota_0 \in \mathbb{W}_2^{\beta}([0,T])$ such that $\iota_0(t) = \mu(t)^{1/2}J_0(t)$ for every $t \in (0,T]$. Then, by DeVore and Lorentz [12], Theorem 2.3 p. 205, there exists a constant $\mathfrak{c}_{\beta,T} > 0$, not depending on m, such that

$$\|p_{\mathcal{S}_m(\mu)}^{\perp}(\mu^{1/2}J_0) - \mu^{1/2}J_0\|_2^2 = \|p_{\mathcal{S}_m(\mu)}^{\perp}(\iota_0) - \iota_0\|_2^2 \leqslant \mathfrak{c}_{\beta,T}m^{-2\beta}.$$

Therefore, by Corollary 3.2,

$$\mathbb{E}(\|\widehat{J}_{m,N} - J_0\|_2^2) \leqslant \|\mu^{-1}\|_{\infty,T} \left(\mathfrak{c}_{\beta,T} m^{-2\beta} + \frac{2m}{N}\right).$$

Finally, consider $m_N \in \{1, ..., N\}$, $\mathcal{M}_N := \{1, ..., m_N\}$ and

$$\widehat{m} = \arg\min_{m \in \mathcal{M}_N} \{ \gamma_{m,N}(\widehat{J}_{m,N}) + \operatorname{pen}(m) \}$$
 with $\operatorname{pen}(.) := \mathfrak{c}_{\operatorname{cal}} \frac{\cdot}{N}$,

where $\mathfrak{c}_{\text{cal}} > 0$ is a constant to calibrate in practice via, for instance, the *slope heuristic*. In the sequel, the φ_j 's fulfill the following assumption.

Assumption 3.3. For every $m, m' \in \{1, ..., N\}$, if m > m', then $S_{m'} \subset S_m$.

The following theorem provides a risk bound on the adaptive estimator $\widehat{J}_{\widehat{m},N}$.

Theorem 3.4. Under Assumptions 2.1 and 3.3, there exists a deterministic constant $\mathfrak{c}_{3.4} > 0$, not depending on N, such that

$$\mathbb{E}(\|\widehat{J}_{\widehat{m},N} - J_0\|_{\langle M \rangle}^2) \leqslant \mathfrak{c}_{3.4} \left(\min_{m \in \mathcal{M}_N} \{ \mathbb{E}(\|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2) + \mathrm{pen}(m) \} + \frac{1}{N} \right).$$

Moreover, under Assumption 2.2,

$$\mathbb{E}(\|\widehat{J}_{\widehat{m},N} - J_0\|_2^2) \leqslant \mathfrak{c}_{3.4} \|\mu^{-1}\|_{\infty,T} \left(\min_{m \in \mathcal{M}_N} \left\{ \|p_{\mathcal{S}_m(\mu)}^{\perp}(\mu^{1/2}J_0) - \mu^{1/2}J_0\|_2^2 + (2 + \mathfrak{c}_{\text{cal}})\frac{m}{N} \right\} + \frac{1}{N} \right).$$

As in the usual nonparametric regression framework, since pen(m) is of same order than the bound on the variance term of $\widehat{J}_{m,N}$ for every $m \in \mathcal{M}_N$, Theorem 3.4 says that the risk of our adaptive estimator is controlled by the minimal risk of $\widehat{J}_{.,N}$ on \mathcal{M}_N up to a multiplicative constant not depending on N.

4. Application to differential equations driven by the fractional Brownian motion

Throughout this section, $V: \mathbb{R} \to \mathbb{R}$ is twice continuously differentiable with bounded derivatives, $\sigma: [0,T] \to \mathbb{R}^*$ is γ -Hölder continuous with $\gamma \in (1-H,1]$, and $b_0: [0,T] \to \mathbb{R}$ is continuous. Under these conditions on V, σ and b_0 , Equation (2) has a unique solution which paths are α -Hölder continuous from [0,T] into \mathbb{R} for every $\alpha \in (1/2,H)$ (see Kubilius et al. [17], Theorem 1.42). The maps V and σ are known and our purpose is to provide a nonparametric estimator of b_0 .

4.1. **Auxiliary model.** The model transformation used in the sequel has been introduced in Kleptsyna and Le Breton [16] in the parametric estimation framework. Consider the function space

$$Q := \{Q : \text{the function } t \mapsto t^{1/2-H}Q(t) \text{ belongs to } \mathbb{L}^1([0,T],dt)\},$$

let $Q_0:[0,T]\to\mathbb{R}$ be the map defined by

$$Q_0(t) := \frac{b_0(t)}{\sigma(t)} ; \forall t \in [0, T],$$

and assume that $Q_0 \in \mathcal{Q}$. Consider also the Molchan martingale M defined by

$$M_t := \int_0^t \ell(t, s) dB_s \; ; \, \forall t \in [0, T],$$

where

$$\ell(t,s) := \mathfrak{c}_H s^{1/2-H} (t-s)^{1/2-H} \mathbf{1}_{(0,t)}(s) \; ; \, \forall s,t \in [0,T]$$

with

$$\mathfrak{c}_H = \left(\frac{\Gamma(3-2H)}{2H\Gamma(3/2-H)^3\Gamma(H+1/2)}\right)^{1/2},$$

and the process Z defined by

$$Z_t := \int_0^t \ell(t, s) dY_s$$
 with $Y_t := \int_0^t \frac{dX_s}{V(X_s)\sigma(s)}$

for every $t \in [0, T]$. Then, Equation (2) leads to

(6)
$$Z_t = j(Q_0)(t) + M_t$$
$$= \int_0^t J(Q_0)(s)d\langle M \rangle_s + M_t,$$

where

$$j(Q)(t) := \int_0^t \ell(t,s)Q(s)ds$$
 and $J(Q)(t) := (2-2H)^{-1}t^{2H-1}j(Q)'(t)$

for every $Q \in \mathcal{Q}$ and almost every $t \in [0,T]$. Note that the Molchan martingale M fulfills Assumption 2.2 with $\mu(t) = t^{1-2H}$ for every $t \in (0,T]$.

4.2. An estimator of Q_0 . In Model (6), for any $m \in \{1, ..., N\}$, the solution $\widehat{J}_{m,N}$ to Problem (4) is a nonparametric estimator of $J(Q_0)$. So, this subsection deals with an estimator of Q_0 solving the inverse problem

$$J(\widehat{Q}) = \widehat{J}_{m,N}.$$

Let us consider the function space

$$\mathcal{J}:=\left\{\iota: \text{the function } t\in[0,T]\longmapsto \int_0^t s^{1-2H}\iota(s)ds \text{ belongs to } \mathcal{I}_{0+}^{3/2-H}(\mathbb{L}^1([0,T],dt))\right\},$$

where $\mathcal{I}_{0+}^{3/2-H}(.)$ is the Riemann-Liouville left-sided fractional integral of order 3/2-H. The reader can refer to Samko et al. [25] on fractional calculus.

In order to provide an estimator of Q_0 with a closed-form expression, let us establish first the following technical proposition.

Proposition 4.1. The map $J: Q \mapsto J(Q)$ is one-to-one from Q into \mathcal{J} . Moreover, for every $\iota \in \mathcal{J}$ and almost every $\iota \in [0,T]$,

$$J^{-1}(\varphi)(t) = \bar{\mathfrak{c}}_H t^{H-1/2} \int_0^t (t-s)^{H-3/2} s^{1-2H} \iota(s) ds$$

with

$$\bar{\mathfrak{c}}_H = \frac{2-2H}{\mathfrak{c}_H \Gamma(3/2-H)\Gamma(H-1/2)}.$$

By Proposition 4.1, if $\varphi_1, \ldots, \varphi_m \in \mathcal{J}$, then

$$\widehat{Q}_{m,N}(t) := \overline{\mathfrak{c}}_H t^{H-1/2} \int_0^t (t-s)^{H-3/2} s^{1-2H} \widehat{J}_{m,N}(s) ds \; ; \; t \in [0,T]$$

is the solution to Problem (7) in \mathcal{Q} . Note that even if the φ_j 's don't belong to \mathcal{J} , since these functions are continuous from [0,T] into \mathbb{R} , $\widehat{Q}_{m,N}$ is well-defined but not necessarily a solution to Problem (7). A simple vector subspace of \mathcal{J} is provided at the end of this subsection.

The following proposition provides risk bounds on $\widehat{Q}_{m,N}$, $m \in \{1,\ldots,N\}$, and on the adaptive estimator $\widehat{Q}_{\widehat{m},N}$.

Proposition 4.2. If the φ_j 's belong to \mathcal{J} , then there exists a deterministic constant $\mathfrak{c}_{4.2,1} > 0$, not depending on N, such that

$$\mathbb{E}(\|\widehat{Q}_{m,N} - Q_0\|_2^2) \leqslant \mathfrak{c}_{4.2,1} \left(\min_{\iota \in \mathcal{S}_m} \|\iota - J(Q_0)\|_{\langle M \rangle}^2 + \frac{m}{N} \right) \; ; \; \forall m \in \{1,\dots,N\}.$$

If in addition the φ_j 's fulfill Assumption 3.3, then there exists a deterministic constant $\mathfrak{c}_{4.2,2} > 0$, not depending on N, such that

$$\mathbb{E}(\|\widehat{Q}_{\widehat{m},N} - Q_0\|_2^2) \leqslant \mathfrak{c}_{4.2,2} \left(\min_{m \in \mathcal{M}_N} \left\{ \min_{\iota \in \mathcal{S}_m} \|\iota - J(Q_0)\|_{\langle M \rangle}^2 + \frac{m}{N} \right\} + \frac{1}{N} \right).$$

Proposition 4.2 says that the MISE of $\widehat{Q}_{m,N}$, $m \in \{1,\ldots,N\}$, (resp. $\widehat{Q}_{\widehat{m},N}$) has at most the same bound than the MISE of $\widehat{J}_{m,N}$ (resp. $\widehat{J}_{\widehat{m},N}$).

Finally, the following proposition provides a simple vector subspace of \mathcal{J} .

Proposition 4.3. The function space

$$\mathbb{J}:=\left\{\iota\in C^1([0,T];\mathbb{R}): \lim_{t\to 0^+}t^{-2H}\iota(t) \ and \ \lim_{t\to 0^+}t^{1-2H}\iota'(t) \ exist \ and \ are \ finite\right\}$$

is a subset of \mathcal{J} .

Consider $m \in \{1, ..., N\}$ and $\psi_1, ..., \psi_m \in C^1([0, T]; \mathbb{R})$ such that $(\psi_1, ..., \psi_m)$ is an orthonormal family of $\mathbb{L}^2([0, T], dt)$. In particular, note that $\psi_1, ..., \psi_m$ are linearly independent. Moreover, assume that $\mathcal{S}_m = \operatorname{span}\{\overline{\varphi}_1, ..., \overline{\varphi}_m\}$ with $\overline{\varphi}_j(t) := t^{2H}\psi_j(t)$ for every $j \in \{1, ..., m\}$ and $t \in [0, T]$. The basis $(\varphi_1, ..., \varphi_m)$ of \mathcal{S}_m , orthonormal in $\mathbb{L}^2([0, T], dt)$, is obtained from $(\overline{\varphi}_1, ..., \overline{\varphi}_m)$ via the Gram-Schmidt process, and the φ_j 's belong to $\mathbb{J} \subset \mathcal{J}$. For every $\iota \in \mathcal{S}_m$, there exist $\alpha_1(\iota), ..., \alpha_m(\iota) \in \mathbb{R}$ such that

$$\iota = \sum_{j=1}^{m} \alpha_j(\iota) \overline{\varphi}_j = \upsilon \mu^{-1/2} \overline{\iota},$$

where $v(t) := t^{H+1/2}$ for every $t \in [0, T]$, and

$$\bar{\iota} := \sum_{j=1}^{m} \alpha_j(\iota) \psi_j \in \mathbb{S}_m \quad \text{with} \quad \mathbb{S}_m := \text{span}\{\psi_1, \dots, \psi_m\}.$$

So, by assuming that $J(Q_0) \in \mathbb{J}$, the bound on the bias term of $\widehat{J}_{m,N}$ in the inequalities of Proposition 4.2 can be controlled the following way:

$$\begin{split} \min_{\iota \in \mathcal{S}_m} \|\iota - J(Q_0)\|_{\langle M \rangle}^2 &= \min_{\overline{\iota} \in \mathbb{S}_m} \|\upsilon(\overline{\iota} - \upsilon^{-1}\mu^{1/2}J(Q_0))\|_2^2 \\ &\leqslant T^{2H+1} \min_{\overline{\iota} \in \mathbb{S}_m} \|\overline{\iota} - \overline{\mu}J(Q_0)\|_2^2 = \|p_{\mathbb{S}_m}^{\perp}(\overline{\mu}J(Q_0)) - \overline{\mu}J(Q_0)\|_2^2 \end{split}$$

where, for every $t \in (0, T]$,

$$\overline{\mu}(t) := v(t)^{-1} \mu(t)^{1/2} = t^{-2H}.$$

If (ψ_1, \ldots, ψ_m) is the *m*-dimensional trigonometric basis, and if there exists $\iota_0 \in \mathbb{W}_2^{\beta}([0,T])$ such that $\iota_0(t) = t^{-2H}J(Q_0)(t)$ for every $t \in (0,T]$, then

$$\min_{\iota \in S} \|\iota - J(Q_0)\|_{\langle M \rangle}^2 \leqslant T^{2H+1} \|p_{\mathbb{S}_m}^{\perp}(\iota_0) - \iota_0\|_2^2 \leqslant \mathfrak{c}_{\beta,T} T^{2H+1} m^{-2\beta}.$$

4.3. Example: drift estimation in a non-autonomous fractional Black-Scholes model. Let us consider a financial market model in which the prices process $S = (S_t)_{t \in [0,T]}$ of the risky asset is defined by

$$S_t := S_0 \exp\left(\int_0^t \left(\overline{b}_0(u) - \frac{\sigma^2}{2} u^{2H-1}\right) du + \sigma B_t\right) \; ; \; \forall t \in [0, T],$$

where $S_0, \sigma \in (0, \infty)$ and $\bar{b}_0 \in C^0([0, T]; \mathbb{R})$. This is a non-autonomous extension of the fractional Black-Scholes model defined in Hu et al. [15]. Thanks to the change of variable formula for Young's integral,

$$S_t = S_0 + \int_0^t S_u \left(\left(\bar{b}_0(u) - \frac{\sigma^2}{2} u^{2H-1} \right) du + \sigma dB_u \right) \; ; \; \forall t \in [0, T].$$

Then, S is the solution to Equation (2) with $V = \mathrm{Id}_{\mathbb{R}}, \, \sigma(.) \equiv \sigma > 0$ and

$$b_0(t) = \overline{b}_0(t) - \frac{\sigma^2}{2}t^{2H-1} ; \forall t \in [0, T].$$

Consider N independent copies S^1, \ldots, S^N of S. For every $i \in \{1, \ldots, N\}$ and $t \in [0, T]$, consider also

$$Z_t^i := \frac{\mathfrak{c}_H}{\sigma} \int_0^t \frac{u^{1/2 - H} (t - u)^{1/2 - H}}{S_u^i} dS_u^i.$$

If the volatility constant σ is known, thanks to Subsection 4.2, a nonparametric estimator of \bar{b}_0 is given by

$$\widehat{\bar{b}}_0(m, N; t) := \frac{\sigma^2}{2} t^{2H-1} + \sigma \widehat{Q}_{m, N}(t) \; ; \; t \in [0, T],$$

where

$$\widehat{Q}_{m,N}(t) := \overline{\mathfrak{c}}_H t^{H-1/2} \int_0^t (t-s)^{H-3/2} s^{1-2H} \widehat{J}_{m,N}(s) ds \; ; \; t \in [0,T]$$

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and

$$\widehat{J}_{m,N} = \arg\min_{J \in \mathcal{S}_m} \left\{ \frac{1}{N} \sum_{i=1}^{N} \left(\int_0^T J(s)^2 s^{1-2H} ds - 2 \int_0^T J(s) dZ_s^i \right) \right\}.$$

Since

$$\|\widehat{\overline{b}}_0(m, N; .) - \overline{b}_0\|_2^2 = \sigma^2 \|\widehat{Q}_{m,N} - Q_0\|_2^2 \text{ with } Q_0 = \frac{\overline{b}_0}{\sigma},$$

if the φ_j 's belong to \mathcal{J} , then Proposition 4.2 provides risk bounds on $\widehat{b}_0(m, N; .)$, $m \in \{1, ..., N\}$, and on the adaptive estimator $\widehat{b}_0(\widehat{m}, N; .)$.

5. Numerical experiments

Some numerical experiments on our estimation method of J_0 in Equation (1) are presented in this subsection when M is the Molchan martingale:

$$M_t = \int_0^t \ell(t, s) dB_s = (2 - 2H)^{1/2} \int_0^t s^{1/2 - H} dW_s \; ; \; t \in [0, 1]$$

with $H \in \{0.6, 0.9\}$ and W the Brownian motion driving the Mandelbrot-Van Ness representation of the fractional Brownian motion B. The estimation method investigated on the theoretical side at Section 3 is implemented here for the three following examples of functions J_0 :

$$J_{0,1}: t \in [0,1] \mapsto 10t^2, \quad J_{0,2}: t \in (0,1] \mapsto 10(-\log(t))^{1/2} \quad \text{and} \quad J_{0,3}: t \in (0,1] \mapsto 20t^{-0.05}.$$

These functions belong to $\mathbb{L}^2([0,1], d\langle M \rangle_t)$ as required. Indeed, one the one hand, $J_{0,1}$ is continuous on [0,1] and

$$\begin{split} -\int_0^1 \log(t) d\langle M \rangle_t &= -\int_0^1 \log(t) t^{1-2H} dt \\ &= \frac{1}{2-2H} \left(\lim_{\varepsilon \to 0^+} \log(\varepsilon) \varepsilon^{2-2H} + \int_0^1 t^{1-2H} dt \right) = \frac{1}{(2-2H)^2} < \infty. \end{split}$$

On the other hand, for every $\alpha \in (0, 1/2)$ such that $H \in (1/2, 1-\alpha)$,

$$\int_0^1 t^{-2\alpha} d\langle M \rangle_t = \int_0^1 t^{1-2\alpha-2H} dt$$

$$= \frac{1}{2(1-\alpha-H)} \left(1 - \lim_{\varepsilon \to 0^+} \varepsilon^{2(1-\alpha-H)} \right) = \frac{1}{2(1-\alpha-H)} < \infty.$$

Since for every $t \in (0,1]$, $J_{0,3}(t) = 20t^{-\alpha}$ with $\alpha = 0.05$, and since $H \in \{0.6,0.9\} \subset (0.5,0.95)$ in our numerical experiments, $J_{0,3}$ is square-integrable with respect to $d\langle M \rangle_t$.

Our adaptive estimator is computed for $J_0 = J_{0,1}$, $J_{0,2}$ or $J_{0,3}$ on N = 300 paths of the process Z observed along the dissection $\{k/n \; ; \; k=1,\ldots,n\}$ of [1/n,1] with n=100, when $(\varphi_1,\ldots,\varphi_m)$ is the m-dimensional trigonometric basis for every $m \in \{1,\ldots,N\}$. This experiment is repeated 100 times, and the means and the standard deviations of the MISE of $\widehat{J}_{\widehat{m},N}$ are stored in Table 1. Moreover, only for H=0.9, 10 estimations (dashed black curves) of $J_{0,1}$, $J_{0,2}$ and $J_{0,3}$ (red curves) are respectively plotted on Figures 1, 2 and 3.

On average, the MISE of our adaptive estimator is higher for the three examples of functions J_0 when H=0.6 than when H=0.9. This is related to the fact that the more the Hurst parameter H is close to 1, the more the paths of B are regular. The standard deviation of the MISE of $\widehat{J}_{\widehat{m},N}$ is also higher when

$J_0; H$	$J_{0,1}; 0.6$	$J_{0,2}; 0.6$	$J_{0,3}; 0.6$	$J_{0,1}; 0.9$	$J_{0,2}; 0.9$	$J_{0,3}; 0.9$
Mean MISE	0.084	0.162	0.079	0.028	0.036	0.022
StD MISE	0.023	0.029	0.022	0.006	0.006	0.005

Table 1. Means and StD of the MISE of $\widehat{J}_{\widehat{m},N}$ (100 repetitions).

H = 0.6.

For both H=0.6 and H=0.9, on average, the MISE of $\widehat{J}_{\widehat{m},N}$ is higher for $J_{0,2}$ than for $J_{0,1}$ and $J_{0,3}$. This is probably related to the fact that $J_{0,2}(t)$ goes faster to infinity when $t\to 0^+$ than $J_{0,3}(t)$, and of course than $J_{0,1}(t)$ which goes not.

Our estimation method seems stable in the sense that for a fixed H (0.6 or 0.9), the standard deviation of the MISE of our adaptive estimator is almost the same for the three examples of functions J_0 . This can be observed on Figures 1, 2 and 3.

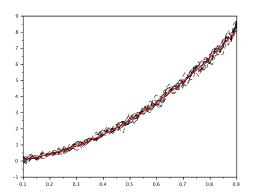


FIGURE 1. Plots of $J_{0,1}$ and of 10 adaptive estimations ($\overline{\widehat{m}} = 51.28$).

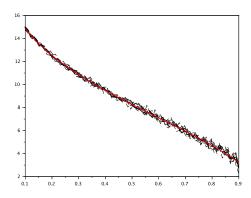


FIGURE 2. Plots of $J_{0,2}$ and of 10 adaptive estimations ($\overline{\hat{m}} = 72.86$).

6. Proofs

6.1. Proof of Proposition 3.1. For every $J, K \in \mathcal{S}_m$,

$$\begin{split} \gamma_{m,N}(J) - \gamma_{m,N}(K) &= \|J\|_{\langle M \rangle}^2 - \|K\|_{\langle M \rangle}^2 - \frac{2}{N} \sum_{i=1}^N \int_0^T (J(s) - K(s)) dZ_s^i \\ &= \|J - J_0\|_{\langle M \rangle}^2 - \|K - J_0\|_{\langle M \rangle}^2 - \frac{2}{N} \sum_{i=1}^N \int_0^T (J(s) - K(s)) dM_s^i. \end{split}$$

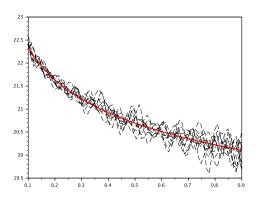


FIGURE 3. Plots of $J_{0,3}$ and of 10 adaptive estimations ($\overline{\widehat{m}} = 39.64$).

Moreover,

$$\gamma_{m,N}(\widehat{J}_{m,N}) \leqslant \gamma_{m,N}(J) ; \forall J \in \mathcal{S}_m.$$

So,

$$\|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2 \le \|J - J_0\|_{\langle M \rangle}^2 + \frac{2}{N} \sum_{i=1}^N \int_0^T (\widehat{J}_{m,N}(s) - J(s)) dM_s^i$$

for any $J \in \mathcal{S}_m$, and then

$$\mathbb{E}(\|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2) \leqslant \|J - J_0\|_{\langle M \rangle}^2 + 2\mathbb{E}\left(\frac{1}{N} \sum_{i=1}^N \int_0^T \widehat{J}_{m,N}(s) dM_s^i\right).$$

Consider $\mathbf{j}_0 = (\langle \varphi_j, J_0 \rangle_{\langle M \rangle})_{j=1,\dots,m}$, and $\mathbf{e} = (\mathbf{e}_1, \dots, \mathbf{e}_m)^*$ such that

$$\mathbf{e}_j := \frac{1}{N} \sum_{i=1}^N \int_0^T \varphi_j(s) dM_s^i \; ; \; \forall j \in \{1, \dots, m\}.$$

Since e is a centered random vector.

$$\begin{split} \mathbb{E}\left(\frac{1}{N}\sum_{i=1}^{N}\int_{0}^{T}\widehat{J}_{m,N}(s)dM_{s}^{i}\right) &= \sum_{j=1}^{m}\mathbb{E}\left(\widehat{\theta}_{j}\cdot\frac{1}{N}\sum_{i=1}^{N}\int_{0}^{T}\varphi_{j}(s)dM_{s}^{i}\right) \\ &= \mathbb{E}(\langle\widehat{\theta},\mathbf{e}\rangle_{2,m}) = \mathbb{E}(\mathbf{e}^{*}\mathbf{\Psi}_{m}^{-1}(\mathbf{j}_{0}+\mathbf{e})) = \mathbb{E}(\mathbf{e}^{*}\mathbf{\Psi}_{m}^{-1}\mathbf{e}). \end{split}$$

Moreover, since M_1, \ldots, M_N are independent copies of M, and since Ψ_m is a symmetric matrix,

$$\mathbb{E}(\mathbf{e}^* \mathbf{\Psi}_m^{-1} \mathbf{e}) = \sum_{j,k=1}^m [\mathbf{\Psi}_m^{-1}]_{j,k} \mathbb{E}(\mathbf{e}_j \mathbf{e}_k) = \frac{1}{N} \sum_{j,k=1}^m [\mathbf{\Psi}_m^{-1}]_{j,k} \int_0^T \varphi_j(s) \varphi_k(s) d\langle M \rangle_s$$
$$= \frac{1}{N} \sum_{k=1}^m \sum_{j=1}^m [\mathbf{\Psi}_m]_{k,j} [\mathbf{\Psi}_m^{-1}]_{j,k} = \frac{1}{N} \sum_{k=1}^m [\mathbf{\Psi}_m \mathbf{\Psi}_m^{-1}]_{k,k} = \frac{m}{N}.$$

Therefore,

$$\mathbb{E}(\|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2) \leqslant \min_{J \in \mathcal{S}_m} \|J - J_0\|_{\langle M \rangle}^2 + \frac{2m}{N}.$$

6.2. **Proof of Corollary 3.2.** Under Assumption 2.2,

$$\min_{J \in \mathcal{S}_m} \|J - J_0\|_{\langle M \rangle}^2 = \min_{J \in \mathcal{S}_m} \|\mu^{1/2} (J - J_0)\|_2^2 = \min_{J \in \mathcal{S}_m(\mu)} \|J - \mu^{1/2} J_0\|_2^2$$

with

$$\mathcal{S}_m(\mu) = \{ \iota \in \mathbb{L}^2([0,T], dt) : \exists \varphi \in \mathcal{S}_m, \forall t \in (0,T], \iota(t) = \mu(t)^{1/2} \varphi(t) \}.$$

Since $S_m(\mu)$ is a closed vector subspace of $\mathbb{L}^2([0,T],dt)$,

(8)
$$\min_{J \in \mathcal{S}_m(\mu)} \|J - \mu^{1/2} J_0\|_2^2 = \|p_{\mathcal{S}_m(\mu)}^{\perp}(\mu^{1/2} J_0) - \mu^{1/2} J_0\|_2^2.$$

Moreover, since $1/\mu$ is continuous from [0,T] into \mathbb{R}_+ under Assumption 2.2,

(9)
$$\|\widehat{J}_{m,N} - J_0\|_2^2 = \|\mu^{-1/2}(\widehat{J}_{m,N} - J_0)\|_{\langle M \rangle}^2 \\ \leqslant \|\mu^{-1}\|_{\infty,T} \|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2.$$

Equality (8) together with Inequality (9) allow to conclude.

6.3. **Proof of Theorem 3.4.** The proof of Theorem 3.4 relies on the following lemma, which is a straightforward consequence of a Bernstein type inequality established in Revuz and Yor [24], p. 153, for continuous local martingales vanishing at 0.

Lemma 6.1. For every $\varepsilon > 0$ and every $\varphi \in \mathbb{L}^2([0,T], d\langle M \rangle_t)$,

$$\mathbb{P}\left(\frac{1}{N}\sum_{i=1}^{N}\int_{0}^{T}\varphi(s)dM_{s}^{i}\geqslant\varepsilon\right)\leqslant\exp\left(-\frac{N\varepsilon^{2}}{2\|\varphi\|_{\langle M\rangle}^{2}}\right).$$

The proof of Theorem 3.4 is dissected in three steps.

Step 1. As established in the proof of Proposition 3.1, for every $J, K \in \mathcal{S}_m$,

$$\gamma_{m,N}(J) - \gamma_{m,N}(K) = \|J - J_0\|_{\langle M \rangle}^2 - \|K - J_0\|_{\langle M \rangle}^2 - \frac{2}{N} \sum_{i=1}^N \int_0^T (J(s) - K(s)) dM_s^i.$$

Moreover,

$$\gamma_{\widehat{m},N}(\widehat{J}_{\widehat{m},N}) + \operatorname{pen}(\widehat{m}) \leqslant \gamma_{m,N}(\widehat{J}_{m,N}) + \operatorname{pen}(m)$$

for any $m \in \mathcal{M}_N$, and then

$$\gamma_{\widehat{m},N}(\widehat{J}_{\widehat{m},N}) - \gamma_{m,N}(\widehat{J}_{m,N}) \leqslant \operatorname{pen}(m) - \operatorname{pen}(\widehat{m}).$$

So, since $2ab \leq a^2 + b^2$ for every $a, b \in \mathbb{R}$, and since $S_m + S_{\widehat{m}} \subset S_{m \vee \widehat{m}}$ under Assumption 3.3,

$$\|\widehat{J}_{\widehat{m},N} - J_0\|_{\langle M \rangle}^2 \leqslant \|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2 + \frac{2}{N} \sum_{i=1}^N \int_0^T (\widehat{J}_{\widehat{m},N}(s) - \widehat{J}_{m,N}(s)) dM_s^i + \operatorname{pen}(m) - \operatorname{pen}(\widehat{m})$$

$$\leqslant \|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2 + \frac{1}{4} \|\widehat{J}_{\widehat{m},N} - \widehat{J}_{m,N}\|_{\langle M \rangle}^2$$

$$+4 \left(\left[\sup_{\varphi \in \mathcal{B}_{m,\widehat{m}}} |\nu_N(\varphi)| \right]^2 - p(m,\widehat{m}) \right)_+ + \operatorname{pen}(m) + 4p(m,\widehat{m}) - \operatorname{pen}(\widehat{m}),$$

where

$$\mathcal{B}_{m,m'} := \{ \varphi \in \mathcal{S}_{m \vee m'} : \|\varphi\|_{\langle M \rangle} \leqslant 1 \} \quad \text{and} \quad p(m,m') := \frac{\mathfrak{c}_{\text{cal}}}{4} \cdot \frac{m \vee m'}{N}$$

for every $m' \in \mathcal{M}_N$, and

$$\nu_N(\varphi) := \frac{1}{N} \sum_{i=1}^N \int_0^T \varphi(s) dM_s^i$$

for every $\varphi \in \mathbb{L}^2([0,T], d\langle M \rangle_t)$. Therefore, since $(a+b)^2 \leq 2a^2 + 2b^2$ for every $a, b \in \mathbb{R}$, and since $4p(m, \widehat{m}) \leq \text{pen}(m) + \text{pen}(\widehat{m})$,

$$(10) \qquad \|\widehat{J}_{\widehat{m},N} - J_0\|_{\langle M \rangle}^2 \leqslant 3\|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2 + 4\mathrm{pen}(m) + 8\left(\left[\sup_{\varphi \in \mathcal{B}_{m,\widehat{m}}} |\nu_N(\varphi)|\right]^2 - p(m,\widehat{m})\right)_{\perp}.$$

Step 2. By using Lemma 6.1, and by following the pattern of the proof of Baraud et al. [1], Proposition 6.1, the purpose of this step is to find a suitable bound on

$$\mathbb{E}\left[\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}|\nu_N(\varphi)|\right]^2-p(m,m')\right)_{\perp}\right]; m'\in\mathcal{M}_N.$$

Consider $\delta_0 \in (0,1)$ and let $(\delta_n)_{n \in \mathbb{N}^*}$ be the real sequence defined by

$$\delta_n := \delta_0 2^{-n} \; ; \, \forall n \in \mathbb{N}^*.$$

Since $\mathcal{S}_{m\vee m'}$ is a vector subspace of $\mathbb{L}^2([0,T],d\langle M\rangle_t)$ of dimension $m\vee m'$, for any $n\in\mathbb{N}$, there exists $T_n\subset\mathcal{B}_{m,m'}$ such that $|T_n|\leqslant (3/\delta_n)^{m\vee m'}$ and, for any $\varphi\in\mathcal{B}_{m,m'}$,

$$\exists f_n \in T_n : \|\varphi - f_n\|_{\langle M \rangle} \leqslant \delta_n$$

In particular, note that

$$\varphi = f_0 + \sum_{n=1}^{\infty} (f_n - f_{n-1}).$$

Then, for any v > 0 and any sequence $(\Delta_n)_{n \in \mathbb{N}}$ of elements of $(0, \infty)$ such that $\Delta := \sum_{n \in \mathbb{N}} \Delta_n < \infty$,

$$\mathbb{P}\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}\nu_{N}(\varphi)\right]^{2} > \upsilon^{2}\Delta^{2}\right)$$

$$= \mathbb{P}\left(\exists (f_{n})_{n\in\mathbb{N}} \in \prod_{n=0}^{\infty} T_{n} : \nu_{N}(f_{0}) + \sum_{n=1}^{\infty} \nu_{N}(f_{n} - f_{n-1}) > \upsilon\Delta\right)$$

$$\leqslant \mathbb{P}\left(\exists (f_{n})_{n\in\mathbb{N}} \in \prod_{n=0}^{\infty} T_{n} : \nu_{N}(f_{0}) > \upsilon\Delta_{0} \text{ or } [\exists n\in\mathbb{N}^{*} : \nu_{N}(f_{n} - f_{n-1}) > \upsilon\Delta_{n}]\right)$$

$$\leqslant \sum_{f_{0}\in T_{0}} \mathbb{P}(\nu_{N}(f_{0}) > \upsilon\Delta_{0}) + \sum_{n=1}^{\infty} \sum_{(f_{n-1}, f_{n})\in\mathbb{T}_{n}} \mathbb{P}(\nu_{N}(f_{n} - f_{n-1}) > \upsilon\Delta_{n})$$

with $\mathbb{T}_n = T_{n-1} \times T_n$ for every $n \in \mathbb{N}^*$. Moreover, $||f_0||_{\langle M \rangle}^2 \leqslant \delta_0^2$ and

$$||f_n - f_{n-1}||_{\langle M \rangle}^2 \le 2\delta_{n-1}^2 + 2\delta_n^2 = \frac{5}{2}\delta_{n-1}^2$$

for every $n \in \mathbb{N}^*$. So, by Lemma 6.1,

$$\mathbb{P}\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}\nu_{N}(\varphi)\right]^{2} > \upsilon^{2}\Delta^{2}\right) \leqslant \sum_{f_{0}\in\mathcal{T}_{0}} \exp\left(-\frac{N\upsilon^{2}\Delta_{0}^{2}}{2\|f_{0}\|_{\langle M\rangle}^{2}}\right) + \sum_{n=1}^{\infty} \sum_{(f_{n-1},f_{n})\in\mathbb{T}_{n}} \exp\left(-\frac{N\upsilon^{2}\Delta_{n}^{2}}{2\|f_{n}-f_{n-1}\|_{\langle M\rangle}^{2}}\right) \\
\leqslant \exp\left(h_{0} - \frac{N\upsilon^{2}\Delta_{0}^{2}}{2\delta_{0}^{2}}\right) + \sum_{n=1}^{\infty} \exp\left(h_{n-1} + h_{n} - \frac{N\upsilon^{2}\Delta_{n}^{2}}{5\delta_{n-1}^{2}}\right)$$
(11)

with $h_n = \log(|T_n|)$ for every $n \in \mathbb{N}$. Now, let us take Δ_0 such that

$$h_0 - \frac{Nv^2\Delta_0^2}{2\delta_0} = -(m \lor m' + x)$$
 with $v, x > 0$,

which leads to

$$\Delta_0 = \left[\frac{2\delta_0^2}{v^2 N} (m \vee m' + x + h_0) \right]^{1/2},$$

and for every $n \in \mathbb{N}^*$, let us take Δ_n such that

$$h_{n-1} + h_n - \frac{Nv^2\Delta_n^2}{5\delta_{n-1}^2} = -(m \lor m' + x + n),$$

which leads to

$$\Delta_n = \left[\frac{5\delta_{n-1}^2}{v^2 N} (m \vee m' + x + h_{n-1} + h_n + n) \right]^{1/2}.$$

For this appropriate sequence $(\Delta_n)_{n\in\mathbb{N}}$,

$$\mathbb{P}\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}\nu_N(\varphi)\right]^2>\upsilon^2\Delta^2\right)\leqslant e^{-x}e^{-(m\vee m')}\left(1+\sum_{n=1}^\infty e^{-n}\right)\leqslant 1.6e^{-x}e^{-(m\vee m')}$$

by Inequality (11), and

$$\Delta^{2} \leqslant \frac{1}{v^{2}N} \left[\sqrt{2}\delta_{0}[(m \vee m' + x)^{1/2} + h_{0}^{1/2}] + \sqrt{5} \sum_{n=1}^{\infty} \delta_{n-1}[(m \vee m' + x)^{1/2} + (h_{n-1} + h_{n} + n)^{1/2}] \right]^{2}$$

$$\leqslant \frac{\delta^{(1)}}{v^{2}N}(m \vee m' + x) + \frac{\delta^{(2)}}{v^{2}N} \leqslant \frac{\delta^{(1)} + \delta^{(2)}}{v^{2}N}(m \vee m' + x)$$

with

$$\delta^{(1)} = 2\left(\sqrt{2}\delta_0 + \sqrt{5}\sum_{n=1}^{\infty}\delta_{n-1}\right)^2 \quad \text{and} \quad \delta^{(2)} = 2\left(\sqrt{2}\delta_0h_0^{1/2} + \sqrt{5}\sum_{n=1}^{\infty}\delta_{n-1}(h_{n-1} + h_n + n)^{1/2}\right)^2.$$

So,

$$\mathbb{P}\left(\left[\sup_{\varphi \in \mathcal{B}_{m,m'}} \nu_N(\varphi)\right]^2 - \frac{\delta^{(1)} + \delta^{(2)}}{v^2} p(m,m') > \frac{\delta^{(1)} + \delta^{(2)}}{v^2 N} x\right) \leqslant 1.6e^{-x} e^{-(m \vee m')}$$

and then, by taking $v = (\delta^{(1)} + \delta^{(2)})^{1/2}$ and y = x/N,

$$\mathbb{P}\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}\nu_N(\varphi)\right]^2-p(m,m')>y\right)\leqslant 1.6e^{-Ny}e^{-(m\vee m')}.$$

Therefore,

$$\mathbb{E}\left[\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}\nu_{N}(\varphi)\right]^{2}-p(m,m')\right)_{+}\right] = \int_{0}^{\infty}\mathbb{P}\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}\nu_{N}(\varphi)\right]^{2}-p(m,m')>y\right)dy$$

$$\leqslant 1.6\frac{e^{-(m\vee m')}}{N}.$$

Step 3. By Inequality (12), there exists a deterministic constant $\mathfrak{c}_1 > 0$, not depending on m and N, such that

$$\mathbb{E}\left[\left(\left[\sup_{\varphi\in\mathcal{B}_{m,\widehat{m}}}|\nu_{N}(\varphi)|\right]^{2}-p(m,\widehat{m})\right)_{+}\right]\leqslant\sum_{m'\in\mathcal{M}_{N}}\mathbb{E}\left[\left(\left[\sup_{\varphi\in\mathcal{B}_{m,m'}}\nu_{N}(\varphi)\right]^{2}-p(m,m')\right)_{+}\right]$$

$$\leqslant\frac{1.6}{N}\sum_{m'\in\mathcal{M}_{N}}e^{-(m\vee m')}\leqslant\frac{1.6}{N}\left(me^{-m}+\sum_{m'>m}e^{-m'}\right)\leqslant\frac{\mathfrak{c}_{1}}{N}.$$

Therefore, by Inequality (10),

$$\mathbb{E}(\|\widehat{J}_{\widehat{m},N} - J_0\|_{\langle M \rangle}^2) \leqslant \min_{m \in \mathcal{M}_N} \{3\mathbb{E}(\|\widehat{J}_{m,N} - J_0\|_{\langle M \rangle}^2) + 4\mathrm{pen}(m)\} + \frac{8\mathfrak{c}_1}{N}.$$

6.4. **Proof of Proposition 4.1.** First of all, recall that

$$Q = \{Q : \text{the function } t \mapsto t^{1/2-H}Q(t) \text{ belongs to } \mathbb{L}^1([0,T],dt)\}$$

and

$$\mathcal{J} = \left\{ \iota : \text{the function } t \in [0, T] \longmapsto \int_0^t s^{1-2H} \iota(s) ds \text{ belongs to } \mathcal{I}_{0+}^{3/2-H}(\mathbb{L}^1([0, T], dt)) \right\}.$$

In the sequel, $\mathcal{I}_{0+}^{\alpha}(.)$ (resp. $\mathcal{D}_{0+}^{\alpha}(.)$) is the Riemann-Liouville left-sided fractional integral (resp. derivative) of order $\alpha > 0$. Consider $Q \in \mathcal{Q}$. Then, the map $Q_H : t \mapsto t^{1/2-H}Q(t)$ belongs to $\mathbb{L}^1([0,T],dt)$,

$$J(Q)(t) = \frac{\mathfrak{c}_H}{2 - 2H} t^{2H - 1} \frac{\partial}{\partial t} \int_0^t s^{1/2 - H} (t - s)^{1/2 - H} Q(s) ds$$
$$= \frac{\mathfrak{c}_H}{2 - 2H} \Gamma(3/2 - H) t^{2H - 1} \mathcal{I}_{0+}^{3/2 - H} (Q_H)'(t)$$

for almost every $t \in [0,T]$, and $J(Q) \in \mathcal{J}$. Therefore, the map J is one-to-one from \mathcal{Q} into \mathcal{J} , and

$$J^{-1}(\iota)(t) = \frac{2 - 2H}{\mathfrak{c}_H \Gamma(3/2 - H)} t^{H - 1/2} \mathcal{D}_{0+}^{3/2 - H} \left(\int_0^{\cdot} s^{1 - 2H} \iota(s) ds \right) (t)$$
$$= \bar{\mathfrak{c}}_H t^{H - 1/2} \int_0^t (t - s)^{H - 3/2} s^{1 - 2H} \iota(s) ds$$

for every $\iota \in \mathcal{J}$ and almost every $t \in [0, T]$.

6.5. **Proof of Proposition 4.2.** Since $\langle M \rangle_t = t^{2-2H}$ for every $t \in [0,T]$ (see Norros et al. [20]), the Molchan martingale M fulfills Assumption 2.2 with $\mu(t) = t^{1-2H}$ for every $t \in (0,T]$. Consider $\iota \in \mathcal{J} \cap \mathbb{L}^2([0,T],d\langle M \rangle_t)$, and note that $J^{-1}(\iota) \in \mathcal{Q}$ thanks to Proposition 4.1. By Cauchy-Schwarz's inequality,

$$||J^{-1}(\iota)||_{2}^{2} = \bar{\mathfrak{c}}_{H}^{2} \int_{0}^{T} t^{2H-1} \left(\int_{0}^{t} (t-s)^{H-3/2} s^{1-2H} \iota(s) ds \right)^{2} dt$$

$$\leqslant \bar{\mathfrak{c}}_{H}^{2} \int_{0}^{T} t^{2H-1} \theta(t) \int_{0}^{t} (t-s)^{H-3/2} s^{1-2H} \iota(s)^{2} ds dt$$

with

$$\theta(t) := \int_0^t (t - s)^{H - 3/2} s^{1 - 2H} ds \; ; \, \forall t \in (0, T].$$

Note that $\overline{\theta}: t \mapsto t^{2H-1}\theta(t)$ is bounded on (0,T]. Indeed, for every $t \in (0,T]$,

$$\begin{split} |\overline{\theta}(t)| &\leqslant t^{2H-1} \left[\left(\frac{t}{2} \right)^{H-3/2} \int_0^{t/2} s^{1-2H} ds + \left(\frac{t}{2} \right)^{1-2H} \int_{t/2}^t (t-s)^{H-3/2} ds \right] \\ &= t^{2H-1} \left[\frac{1}{2-2H} \left(\frac{t}{2} \right)^{H-3/2} \left(\frac{t}{2} \right)^{2-2H} + \frac{1}{H-1/2} \left(\frac{t}{2} \right)^{1-2H} \left(\frac{t}{2} \right)^{H-1/2} \right] \\ &= \frac{1}{2^{1/2-H}} \left(\frac{1}{2-2H} + \frac{1}{H-1/2} \right) t^{H-1/2} \xrightarrow[t \to 0^+]{} 0. \end{split}$$

So, by Fubini's theorem,

$$||J^{-1}(\iota)||_{2}^{2} \leqslant \overline{\mathfrak{c}}_{H}^{2} ||\overline{\theta}||_{\infty,T} \int_{0}^{T} \int_{0}^{T} (t-s)^{H-3/2} s^{1-2H} \iota(s)^{2} \mathbf{1}_{(s,T]}(t) ds dt$$

$$= \overline{\mathfrak{c}}_{H}^{2} ||\overline{\theta}||_{\infty,T} \int_{0}^{T} s^{1-2H} \iota(s)^{2} \int_{s}^{T} (t-s)^{H-3/2} dt ds$$

$$\leqslant (H-1/2)^{-1} \overline{\mathfrak{c}}_{H}^{2} T^{H-1/2} ||\overline{\theta}||_{\infty,T} \int_{0}^{T} s^{1-2H} \iota(s)^{2} ds.$$

Then, $J^{-1}(\iota) \in \mathcal{Q} \cap \mathbb{L}^2([0,T],dt)$ and

(13)
$$||J^{-1}(\iota)||_{2}^{2} \leqslant \mathfrak{c}_{1} ||\iota||_{\langle M \rangle}^{2} \quad \text{with} \quad \mathfrak{c}_{1} = (H - 1/2)^{-1} \overline{\mathfrak{c}}_{H}^{2} T^{H - 1/2} ||\overline{\theta}||_{\infty, T}.$$

The conclusion comes from Inequality (13), Proposition 3.1 and Theorem 3.4.

6.6. Proof of Proposition 4.3. Thanks to Samko et al. [25], Theorem 2.3,

$$\mathcal{J} = \left\{ \iota : \mathcal{I}_{0+}^{H-1/2} \left(\int_0^{\cdot} s^{1-2H} \iota(s) ds \right) \text{ is absolutely continuous from } [0,T] \text{ into } \mathbb{R} \right\}.$$

So, since Lipschitz continuous functions are absolutely continuous,

$$\mathcal{J}_{\text{Lip}} := \left\{ \iota : \mathcal{I}_{0+}^{H-1/2} \left(\int_0^{\cdot} s^{1-2H} \iota(s) ds \right) \text{ is Lipschitz continuous from } [0, T] \text{ into } \mathbb{R} \right\}$$

is a subset of \mathcal{J} . Moreover, for every $\iota \in \mathcal{J}_{Lip}$ and $t \in [0,T]$, by Fubini's theorem.

$$\begin{split} \mathcal{I}_{0+}^{H-1/2} \left(\int_0^{\cdot} s^{1-2H} \iota(s) ds \right)(t) &= \frac{1}{\Gamma(H-1/2)} \int_0^t (t-u)^{H-3/2} \int_0^u s^{1-2H} \iota(s) ds du \\ &= \frac{1}{\Gamma(H-1/2)} \int_0^t s^{1-2H} \iota(s) \int_s^t (t-u)^{H-3/2} du ds \\ &= \frac{1}{(H-1/2)\Gamma(H-1/2)} \int_0^t (t-s)^{H-1/2} s^{1-2H} \iota(s) ds. \end{split}$$

Then,

$$\mathcal{J}_{\mathrm{Lip}} = \left\{ \iota : \text{the function } t \in [0,T] \longmapsto \int_0^t (t-s)^{H-1/2} s^{1-2H} \iota(s) ds \text{ is Lipschitz continuous} \right\}$$

Now, recall that

$$\mathbb{J} = \left\{ \iota \in C^1([0,T];\mathbb{R}) : \lim_{t \to 0^+} t^{-2H} \iota(t) \text{ and } \lim_{t \to 0^+} t^{1-2H} \iota'(t) \text{ exist and are finite} \right\}$$

and let us show that $\mathbb{J} \subset \mathcal{J}_{\text{Lip}} \subset \mathcal{J}$. For any $\iota \in \mathbb{J}$ and every $t \in [0, T]$,

$$\begin{split} \bar{\iota}(t) &:= \int_0^t (t-s)^{H-1/2} s^{1-2H} \iota(s) ds \\ &= -\lim_{\varepsilon \to 0} \left[s^{1-2H} \iota(s) \frac{(t-s)^{H+1/2}}{H+1/2} \right]_\varepsilon^t \\ &+ \int_0^t ((1-2H) s^{-2H} \iota(s) + s^{1-2H} \iota'(s)) \frac{(t-s)^{H+1/2}}{H+1/2} ds = \int_0^t \theta(s) (t-s)^{H+1/2} ds \end{split}$$

with

$$\theta(s) := \frac{1}{H + 1/2} ((1 - 2H)s^{-2H}\iota(s) + s^{1-2H}\iota'(s)) \; ; \; \forall s \in (0, T].$$

Since $\iota \in \mathbb{J}$, the map $\theta : t \mapsto \theta(t)$ is bounded on (0,T], and for every $t_1, t_2 \in [0,T]$ such that $t_2 > t_1$,

$$|\bar{\iota}(t_2) - \bar{\iota}(t_1)| \leq \int_{t_1}^{t_2} \theta(s)(t_2 - s)^{H+1/2} ds + \int_{0}^{t_1} \theta(s)|(t_2 - s)^{H+1/2} - (t_1 - s)^{H+1/2}|ds| \leq (H + 3/2)\|\theta\|_{\infty, T} T^{H+1/2} |t_2 - t_1|.$$

So, the function $t \mapsto \bar{\iota}(t)$ is Lipschitz continuous from [0,T] into \mathbb{R} , which leads to $\iota \in \mathcal{J}_{\text{Lip}}$.

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