



HAL
open science

Stochastic elliptic operators defined by non-Gaussian random fields with uncertain spectrum

Christian Soize

► **To cite this version:**

Christian Soize. Stochastic elliptic operators defined by non-Gaussian random fields with uncertain spectrum. *Theory of Probability and Mathematical Statistics*, 2021, 105, pp.113-136. 10.1090/tpms/1159 . hal-03381362

HAL Id: hal-03381362

<https://hal.science/hal-03381362>

Submitted on 14 Dec 2021

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

STOCHASTIC ELLIPTIC OPERATORS DEFINED BY NON-GAUSSIAN RANDOM FIELDS WITH UNCERTAIN SPECTRUM

CHRISTIAN SOIZE

ABSTRACT. This paper presents a construction and the analysis of a class of non-Gaussian positive-definite matrix-valued homogeneous random fields with uncertain spectral measure for stochastic elliptic operators. Then the stochastic elliptic boundary value problem in a bounded domain of the 3D-space is introduced and analyzed for stochastic homogenization.

1. INTRODUCTION

Random fields theory has extensively been developed [35, 25, 1, 34, 24], in particular in the context of continuum physics [19, 32, 17]. The framework of this paper is that of the analysis of the stochastic homogenization of a 3D-linear anisotropic elastic random medium. The elasticity field is modeled by a Non-Gaussian positive-definite fourth-order tensor-valued homogeneous random field. This paper present an extension of the works [29, 18, 31, 32] devoted to random field representations for stochastic elliptic boundary value problems and stochastic homogenization. We propose a novel probabilistic modeling to take into account uncertainties in the spectral measure of the elasticity random field and we analyze the stochastic elliptic boundary value problem (BVP) that has to be solved to perform the stochastic homogenization.

Notations

The following notations are used:

x : lower-case Latin or Greek letters are deterministic real variables.

\mathbf{x} : boldface lower-case Latin, Greek, and calligraphic letters are deterministic vectors.

X : upper-case Latin, Greek, and calligraphic letters are real-valued random variables.

\mathbf{X} : boldface upper-case Latin or Greek letters are vector-valued random variables.

$[x]$: lower-case Latin or Greek letters between brackets are deterministic matrices.

$[\mathbf{X}]$: boldface upper-case letters between brackets are matrix-valued random variables.

\mathbb{C} : fourth-order tensor-valued random field.

\mathbb{N}, \mathbb{R} : set of all the integers $\{0, 1, 2, \dots\}$, set of all the real numbers.

\mathbb{R}^n : Euclidean vector space on \mathbb{R} of dimension n .

$\mathbb{M}_{n,m}$: set of all the $(n \times m)$ real matrices.

\mathbb{M}_n : set of all the square $(n \times n)$ real matrices.

\mathbb{M}_n^S : set of all the symmetric $(n \times n)$ real matrices.

\mathbb{M}_n^+ : set of all the positive-definite symmetric $(n \times n)$ real matrices.

$[I_n]$: identity matrix in \mathbb{M}_n .

$\mathbf{x} = (x_1, \dots, x_n)$: point in \mathbb{R}^n .

1991 *Mathematics Subject Classification.* Primary 60G60, 35J25; Secondary 74Q05, 74A40.

Key words and phrases. Non-Gaussian random field, uncertain spectral measure, stochastic elliptic boundary value problem, stochastic homogenization, random effective elasticity tensor.

- $\langle \mathbf{x}, \mathbf{y} \rangle_2 = x_1 y_1 + \dots + x_n y_n$: inner product in \mathbb{R}^n .
 $\|\mathbf{x}\|_2$: norm in \mathbb{R}^n such that $\|\mathbf{x}\|^2 = \langle \mathbf{x}, \mathbf{x} \rangle_2$.
 $\| [a] \|_2 = \sup_{\mathbf{x}} \{ \|[a]\mathbf{x}\|_2 / \|\mathbf{x}\|_2 \}$ for $[a] \in \mathbb{M}_n$ and $\mathbf{x} \in \mathbb{R}^n$.
 $[x]^T$: transpose of matrix $[x]$.
 $\text{tr}\{[x]\}$: trace of the square matrix $[x]$.
 $\langle [x], [y] \rangle_F = \text{tr}\{[x]^T [y]\}$, inner product of matrices $[x]$ and $[y]$ in $\mathbb{M}_{n,m}$.
 $\|[x]\|_F$: Frobenius norm of matrix $[x]$ such that $\|[x]\|_F^2 = \langle [x], [x] \rangle_F$.
 $\mathbb{1}_B$: indicatrix function of set B .
 i : imaginary unit.
 $\delta_{kk'}$: Kronecker's symbol.
 $\delta_{\mathbf{x}_0}$: Dirac measure at point \mathbf{x}_0 .
 $a.s.$: almost surely.
 E : mathematical expectation.
- $\mathcal{C}_w, \mathcal{C}_y, \mathcal{C}_z, \mathcal{C}_\xi, \mathcal{C}_\varphi, \mathcal{C}_\mathcal{S}$: set of values for $\mathbf{w}, [y], z, \xi, \varphi, \mathcal{S}$.
 δ_c : controls the level of statistical fluctuations in the random medium.
 δ_s : controls the level of uncertainties in the uncertain spectral measure.
 G : Gaussian random field indexed by \mathbb{R}^3 .
 G^ν : normalized random field corresponding to a finite representation of G .
 \mathbf{k} : wave vector in \mathbb{R}^3 .
 $\bar{\mathbb{K}}$: compact support of s.d.f $\mathbf{k} \mapsto s(\mathbf{k})$.
 s : spectral density function (s.d.f) $\mathbf{k} \mapsto s(\mathbf{k})$ of random field G .
 \mathcal{S} : parameter $\{\mathbf{w}, [y]\}$ controlling the uncertainties in random field G^ν .
 \mathbf{S} : random variable $\{\mathbf{W}, [\mathbf{Y}]\}$ modeling $\mathcal{S} = \{\mathbf{w}, [y]\}$.
 $\underline{\mathcal{S}}$: mean value $\{\underline{\mathbf{w}}, [y]\}$ of $\mathcal{S} = \{\mathbf{W}, [\mathbf{Y}]\}$.
 \mathbf{w} : in $\mathcal{C}_w \subset \mathbb{R}^3$, controls the compact support $\bar{\mathbb{K}}$ of s.d.f $\mathbf{k} \mapsto s(\mathbf{k})$.
 $\underline{\mathbf{w}}$: mean value of \mathbf{W} .
 \mathbf{W} : random vector modelling \mathbf{w} .
 χ : dimensionless s.d.f $\tau \mapsto \chi(\tau)$ on \mathbb{R}^3 with compact support $[-1, 1]^3$.
 $[y]$: in $\mathcal{C}_y \subset \mathbb{M}_{3, \hat{\nu}_s}$, controls the dimensionless s.d.f $\tau \mapsto \chi(\tau)$.
 $\underline{[y]}$: mean value of $[\mathbf{Y}]$.
 $[\mathbf{Y}]$: random vector modelling $[y]$.

2. NON-GAUSSIAN RANDOM FIELD WITH UNCERTAIN SPECTRAL MEASURE

Physical framework of the considered random fields class. For stochastic homogenization of linear elastic heterogeneous media presented in Section 5, the physical space \mathbb{R}^3 is referred to a Cartesian reference system for which the generic point is $\mathbf{x} = (x_1, x_2, x_3)$. Nevertheless, all the developments presented in Sections 2 to 4, can easily be adapted to any finite dimension greater or equal to 1. We consider a linear elastic heterogeneous medium for which the elasticity field is a non-Gaussian fourth-order tensor-valued random field $\tilde{\mathbb{C}} = \{\tilde{\mathbb{C}}_{ijpq}\}_{ijpq}$ with $i, j, p,$ and q in $\{1, 2, 3\}$. A general probabilistic construction has been proposed in [9, 10, 32] in order to take into account the material symmetry in a given symmetry class for the mean value of the elasticity random field and considering the statistical fluctuations either in the same symmetry class, or in another symmetry class, or in a mixture of two symmetry classes. In this paper, we start the construction of the random field with the initial formulation proposed in [29]. It is thus assumed that the mean value of the elasticity random field $\underline{\mathbb{C}}$ is independent of \mathbf{x} and belongs to any symmetry class. The statistical fluctuations of $\tilde{\mathbb{C}}$ around $\underline{\mathbb{C}}$ are assumed

to be anisotropic and statistically homogeneous in \mathbb{R}^3 (it should be noted that the developments presented could be extended to a more general case of material symmetry for the statistical fluctuations but would greatly complicate the presentation). An important quantity that controls the statistical fluctuations is the spectral measure that allows the spatial correlation structure to be described (see for instance [27, 12, 34, 14, 16]) and that we will assumed to be uncertain in this paper.

Principle of construction of the uncertain spectral measure. The uncertainties are modeled using the probability theory. A parameterization of the spectral measure, involving a parameter \mathcal{S} , is introduced. The uncertain spectral measure is obtained by modeling \mathcal{S} by a random variable \mathbf{S} . We then construct a non-Gaussian positive-definite fourth-order tensor-valued homogeneous random field $\mathbb{C}(\cdot; \mathcal{S})$ parameterized with \mathcal{S} such that $\tilde{\mathbb{C}} = \mathbb{C}(\cdot; \mathbf{S})$. Throughout the paper the quantities surmounted by a tilde correspond to the case of uncertain spectral measure modeled by a random spectral measure.

Non-Gaussian positive-definite matrix-valued homogeneous random fields with uncertain spectral measure. For all \mathbf{x} fixed in \mathbb{R}^3 , the fourth-order random tensor $\mathbb{C}(\mathbf{x}; \mathcal{S})$ will verify the usual properties: symmetry, positivity, and existence of a positive lower bound. Let $\mathbf{i} = (i, j)$ with $1 \leq i \leq j \leq 3$ and $\mathbf{j} = (p, q)$ with $1 \leq p \leq q \leq 3$ be the indices with values in $\{1, \dots, 6\}$, which allow for defining the \mathbb{M}_6^+ -valued random matrix $[\mathbb{C}(\mathbf{x}; \mathcal{S})]$ such that $[\mathbb{C}(\mathbf{x}; \mathcal{S})]_{\mathbf{ij}} = \mathbb{C}_{ijpq}(\mathbf{x}; \mathcal{S})$ (use of the representation in Voigt notation for the constitutive equation).

Random effective elasticity matrix. For fixed \mathcal{S} , the parameterized effective elasticity matrix $[\mathbb{C}^{\text{eff}}(\mathcal{S})]$ is a random matrix in \mathbb{M}_6^+ , which is obtained by stochastic homogenization solving a stochastic elliptic BVP on a bounded domain Ω of \mathbb{R}^3 . The random effective elasticity matrix $[\tilde{\mathbb{C}}^{\text{eff}}]$, corresponding to the elasticity random field with uncertain spectral measure, is then given by $[\tilde{\mathbb{C}}^{\text{eff}}] = [\mathbb{C}^{\text{eff}}(\mathbf{S})]$.

Definition 2.1 (*Non-Gaussian homogeneous random field $[\mathbb{C}(\cdot; \mathcal{S})]$ given \mathcal{S}*). Let $[\underline{\mathbb{C}}]$ be a given matrix in \mathbb{M}_6^+ independent of \mathbf{x} and \mathcal{S} . We define $\{[\mathbb{C}(\mathbf{x}; \mathcal{S})], \mathbf{x} \in \mathbb{R}^3\}$ as a non-Gaussian \mathbb{M}_6^+ -valued second-order random field, on a probability space $(\Theta, \mathcal{T}, \mathcal{P})$, indexed by \mathbb{R}^3 , homogeneous, mean-square continuous, whose mean value is $[\underline{\mathbb{C}}] = E\{[\mathbb{C}(\mathbf{x}; \mathcal{S})]\}$ that is therefore independent of \mathbf{x} and \mathcal{S} . We have

$$(2.1) \quad \text{tr}[\underline{\mathbb{C}}] = \underline{c}_1 \quad , \quad \langle [\underline{\mathbb{C}}]\boldsymbol{\omega}, \boldsymbol{\omega} \rangle_2 \geq \underline{c}_0 \|\boldsymbol{\omega}\|_2^2, \quad \forall \boldsymbol{\omega} \in \mathbb{R}^6,$$

in which \underline{c}_0 and \underline{c}_1 are two positive finite constants.

With the construction proposed in this paper, $E\{[\tilde{\mathbb{C}}(\mathbf{x})]\} = E\{[\mathbb{C}(\mathbf{x}; \mathcal{S})]\}$ will not be equal to $[\underline{\mathbb{C}}]$ (that is not a difficulty). However, we will see that $E\{[\tilde{\mathbb{C}}(\mathbf{x})]\} \simeq [\underline{\mathbb{C}}]$.

Lemma 2.2 (*Normalization of random field $[\mathbb{C}(\cdot; \mathcal{S})]$ given \mathcal{S}*). Let $[\underline{\mathbb{L}}]$ be the upper triangular (6×6) real matrix such that $[\underline{\mathbb{C}}] = [\underline{\mathbb{L}}]^T [\underline{\mathbb{L}}]$. For fixed \mathcal{S} , the normalized representation of $[\mathbb{C}(\mathbf{x}; \mathcal{S})]$ is written as,

$$(2.2) \quad [\mathbb{C}(\mathbf{x}; \mathcal{S})] = \frac{1}{1 + \epsilon} [\underline{\mathbb{L}}]^T (\epsilon [I_6] + [\mathbb{C}(\mathbf{x}; \mathcal{S})]) [\underline{\mathbb{L}}],$$

in which $\epsilon > 0$ is given and where $\{[\mathbb{C}(\mathbf{x}; \mathcal{S})], \mathbf{x} \in \mathbb{R}^3\}$ is a \mathbb{M}_6^+ -valued random field (by construction), defined on $(\Theta, \mathcal{T}, \mathcal{P})$, indexed by \mathbb{R}^3 . Then $[\mathbb{C}(\cdot; \mathcal{S})]$ is homogeneous, mean-square continuous, and such that

$$(2.3) \quad E\{[\mathbb{C}(\mathbf{x}; \mathcal{S})]\} = [I_6] \quad , \quad \forall \mathbf{x} \in \mathbb{R}^3.$$

Proof. (Lemma 2.2). Under the hypotheses introduced in Definition 2.1, it is easy to see that $[\mathbf{C}(\cdot; \mathbf{S})]$ is a second-order, homogeneous, mean-square continuous random field, and satisfies Eq. (2.3). \square

It should be noted that the lower bound $\epsilon [\mathbb{C}]/(1+\epsilon)$ used in Eq. (2.2) could be replaced by a more general lower bound $[\mathbb{C}_b] \in \mathbb{M}_6^+$ as proposed in [32, 31]. Note also that, as previously, introducing $[\tilde{\mathbf{C}}(\mathbf{x})] = [\mathbf{C}(\mathbf{x}; \mathbf{S})]$, $E\{[\tilde{\mathbf{C}}(\mathbf{x})] \}$ will not be equal to $[I_6]$ and we will see that $E\{[\tilde{\mathbf{C}}(\mathbf{x})]\} \simeq [I_6]$.

Hypothesis 1 (*Principle of construction of random field $[\mathbb{C}(\cdot; \mathbf{S})]$ given \mathbf{S}*). By construction (see Lemma 2.2), $[\mathbb{C}(\cdot; \mathbf{S})]$ is a \mathbb{M}_6^+ -valued random field indexed by \mathbb{R}^3 and homogeneous. For \mathbf{x} fixed in \mathbb{R}^3 , the \mathbb{M}_6^+ -valued random variable $[\mathbf{C}(\mathbf{x}; \mathbf{S})]$ is constructed by using the Maximum Entropy Principle under the following available information,

$$(2.4) \quad E\{[\mathbf{C}(\mathbf{x}; \mathbf{S})]\} = [I_6] \quad , \quad E\{\log(\det[\mathbf{C}(\mathbf{x}; \mathbf{S})])\} = b_c ,$$

in which b_c is independent of \mathbf{x} and \mathbf{S} and such that $|b_c| < +\infty$. The second equality is introduced in order that the random matrix $[\mathbf{C}(\mathbf{x}; \mathbf{S})]^{-1}$ (that exists almost surely) be a second-order random variable: $E\{\|[\mathbf{C}(\mathbf{x}; \mathbf{S})]^{-1}\|_F^2\} \leq E\{\|[\mathbf{C}(\mathbf{x}; \mathbf{S})]^{-1}\|_F^2\} < +\infty$. With such a construction, $[\mathbf{C}(\mathbf{x}; \mathbf{S})]$ will appear as a nonlinear transformation of $6 \times (6+1)/2 = 21$ independent normalized Gaussian real-valued random variables $\{G_{mn}(\mathbf{x}; \mathbf{S}), 1 \leq m \leq n \leq 6\}$, such that

$$(2.5) \quad E\{G_{mn}(\mathbf{x}; \mathbf{S})\} = 0 \quad , \quad E\{G_{mn}(\mathbf{x}; \mathbf{S})^2\} = 1 .$$

The spatial correlation structure of random field $\{[\mathbf{C}(\mathbf{x}; \mathbf{S})], \mathbf{x} \in \mathbb{R}^3\}$ is introduced by considering 21 independent real-valued random fields $\{G_{mn}(\mathbf{x}; \mathbf{S}), \mathbf{x} \in \mathbb{R}^3\}$ for $1 \leq m \leq n \leq 6$, corresponding to 21 independent copies of a unique normalized Gaussian homogeneous mean-square continuous real-valued random field $\{G(\mathbf{x}; \mathbf{S}), \mathbf{x} \in \mathbb{R}^3\}$ given its normalized spectral measure parameterized by \mathbf{S} . Note that the Gaussian random field $G(\cdot; \mathbf{S})$ is entirely defined by its normalized spectral measure (parameterized by \mathbf{S}) because, $\forall \mathbf{x} \in \mathbb{R}^3$, $E\{G(\mathbf{x}; \mathbf{S})\} = 0$ and $E\{G(\mathbf{x}; \mathbf{S})^2\} = 1$. The constant b_c is eliminated in favor of a hyperparameter δ_c , which allows for controlling the level of statistical fluctuations of $[\mathbf{C}(\mathbf{x}; \mathbf{S})]$, defined by $\delta_c = (E\{\|[\mathbf{C}(\mathbf{x}; \mathbf{S})] - [I_6]\|_F^2\}/6)^{1/2}$, independent of \mathbf{x} and chosen independent of \mathbf{S} .

Proposition 2.1 (Random field $[\mathbf{C}(\cdot; \mathbf{S})]$). *Let us assume Hypothesis 1.*

(i) *Let $d^{\mathbf{S}}C = 2^{15/2} \Pi_{1 \leq m \leq n \leq 6} dC_{mn}$ be the volume element on Euclidean space $\mathbb{M}_6^{\mathbf{S}}$ in which dC_{mn} is the Lebesgue measure on \mathbb{R} . For all \mathbf{x} fixed in \mathbb{R}^3 , the probability measure $P_{[\mathbf{C}(\mathbf{x}; \mathbf{S})]}(d^{\mathbf{S}}C)$ of the \mathbb{M}_6^+ -valued random variable $[\mathbf{C}(\mathbf{x}; \mathbf{S})]$ constructed with the Maximum Entropy Principle under the constraints defined by Eq. (2.4), is independent of \mathbf{x} (homogeneous random field), independent of \mathbf{S} (this marginal probability measure does not depend of the correlation structure), and written as $P_{[\mathbf{C}(\mathbf{x}; \mathbf{S})]}(d^{\mathbf{S}}C) = p_{[\mathbf{C}(\mathbf{x}; \mathbf{S})]}([C]) d^{\mathbf{S}}C$ in which the probability density function is written as $p_{[\mathbf{C}(\mathbf{x}; \mathbf{S})]}([C]) = \mathbb{1}_{\mathbb{M}_6^+}([C]) c_c (\det[C])^{7(1-\delta_c^2)/(2\delta_c^2)} \exp(-7\text{tr}\{[C]\}/(2\delta_c^2))$ with c_c the normalization constant and where hyperparameter δ_c must belong to the real interval $]0, \sqrt{7/11}[$.*

(ii) *For all \mathbf{x} fixed in \mathbb{R}^3 , random matrix $[\mathbf{C}(\mathbf{x}; \mathbf{S})]$ is written as*

$$(2.6) \quad [\mathbf{C}(\mathbf{x}; \mathbf{S})] = [\mathbf{L}(\mathbf{x}; \mathbf{S})]^T [\mathbf{L}(\mathbf{x}; \mathbf{S})] ,$$

in which $[\mathbf{L}(\mathbf{x}; \mathbf{S})]$ is an upper triangular random matrix in \mathbb{M}_6 such that

- 1) the 21 random variables $\{[\mathbf{L}(\mathbf{x}; \mathbf{S})]_{mn}, 1 \leq m \leq n \leq 6\}$ are mutually independent.
- 2) for $1 \leq m < n \leq 6$, $[\mathbf{L}(\mathbf{x}; \mathbf{S})]_{mn} = \sigma_c G_{mn}(\mathbf{x}; \mathbf{S})$ with $\sigma_c = \delta_c/\sqrt{7}$ and where $G_{mn}(\mathbf{x}; \mathbf{S})$ is a normalized Gaussian real-valued random variable (see Eq. (2.5)).

3) for $1 \leq m = n \leq 6$, $[\mathbf{L}(\mathbf{x}; \mathbf{S})]_{mm} = \sigma_c \sqrt{2h(G_{mm}(\mathbf{x}; \mathbf{S}); \alpha_m)}$ with $\alpha_m = 1/(2\sigma_c^2) + (1-m)/2$ such that $\alpha_1 > \dots > \alpha_6 > 3$ and where $G_{mm}(\mathbf{x}; \mathbf{S})$ is a normalized Gaussian real-valued random variable (see Eq. (2.5)). The function $b \mapsto h(b; \alpha)$ is such that $\mathbb{F}_\alpha = h(\mathcal{N}; \alpha)$ is a Gamma random variable with parameter α when \mathcal{N} is the normalized Gaussian real-valued random variable.

(iii) The 21 random fields $\{G_{mn}(\mathbf{x}; \mathbf{S}), \mathbf{x} \in \mathbb{R}^3\}$ for $1 \leq m \leq n \leq 6$ are 21 independent copies of a normalized Gaussian homogeneous mean-square continuous real-valued random field $\{G(\mathbf{x}; \mathbf{S}), \mathbf{x} \in \mathbb{R}^3\}$,

$$(2.7) \quad E\{G(\mathbf{x}; \mathbf{S})\} = 0 \quad , \quad E\{G(\mathbf{x}; \mathbf{S})^2\} = 1 \quad , \quad \forall \mathbf{x} \in \mathbb{R}^3,$$

and which will be defined in Section 3 for imposing its spatial correlation structure via its spectral measure. parameterized by \mathbf{S} .

Proof. (Proposition 2.1). We refer the reader to [28] for the construction using the Maximum Entropy Principle and to [29, 31, 32] for the representation defined by Eq. (2.6). However, we have to prove the properties that yield $E\{[\mathbf{C}(\mathbf{x}; \mathbf{S})]\} = [I_6]$, this proof being used in Remark 2.4. For $1 \leq m \leq n \leq 6$, $[\mathbf{C}(\mathbf{x}; \mathbf{S})]_{mn} = \sum_{\ell=1}^m [\mathbf{L}(\mathbf{x}; \mathbf{S})]_{\ell m} [\mathbf{L}(\mathbf{x}; \mathbf{S})]_{\ell n}$. For $m = n$, $E\{[\mathbf{C}(\mathbf{x}; \mathbf{S})]_{mm}\} = 2\sigma_c^2 E\{h(G_{mm}(\mathbf{x}; \mathbf{S}); \alpha_m)\} + \sigma_c^2 \sum_{\ell < m} E\{G_{\ell m}(\mathbf{x}; \mathbf{S})^2\}$. Equation (2.5) yields $\sum_{\ell < m} E\{G_{\ell m}(\mathbf{x}; \mathbf{S})^2\} = \sum_{\ell < m} 1 = m-1$ and $E\{h(G_{mm}(\mathbf{x}; \mathbf{S}); \alpha_m)\} = E\{\mathbb{F}_{\alpha_m}\} = \alpha_m = 1/(2\sigma_c^2) + (1-m)/2$. Therefore, $E\{[\mathbf{C}(\mathbf{x}; \mathbf{S})]_{mm}\} = 1$. For $1 \leq m < n \leq 6$, we have $E\{[\mathbf{C}(\mathbf{x}; \mathbf{S})]_{mn}\} = \sigma_c^2 E\{G_{mn}(\mathbf{x}; \mathbf{S}) \sqrt{2h(G_{mm}(\mathbf{x}; \mathbf{S}); \alpha_m)}\} + \sigma_c^2 \sum_{\ell < m} E\{G_{\ell m}(\mathbf{x}; \mathbf{S}) G_{\ell n}(\mathbf{x}; \mathbf{S})\}$. Proposition 2.1-(iii) and Eq. (2.5) yield $E\{G_{\ell m}(\mathbf{x}; \mathbf{S}) G_{\ell n}(\mathbf{x}; \mathbf{S})\} = E\{G_{\ell m}(\mathbf{x}; \mathbf{S})\} \times E\{G_{\ell n}(\mathbf{x}; \mathbf{S})\} = 0$ and $E\{G_{mn}(\mathbf{x}; \mathbf{S}) \sqrt{2h(G_{mm}(\mathbf{x}; \mathbf{S}); \alpha_m)}\} = E\{G_{mn}(\mathbf{x}; \mathbf{S})\} \times E\{\sqrt{2h(G_{mm}(\mathbf{x}; \mathbf{S}); \alpha_m)}\} = 0$. Therefore, $E\{[\mathbf{C}(\mathbf{x}; \mathbf{S})]_{mn}\} = 0$. We thus obtain the first equation in Eq. (2.4). \square

Lemma 2.3 (Properties of function h). (i) Let be $\alpha > 3$. Function $b \mapsto h(b; \alpha) : \mathbb{R} \rightarrow]0, +\infty[$ defined in Proposition 2.1 is written as $h(b; \alpha) = F_\alpha^{-1}(F(b))$ in which $F(b) = \int_{-\infty}^b (2\pi)^{-1/2} e^{-t^2/2} dt$ and where F_α^{-1} is the reciprocal function of $F_\alpha(\mathfrak{g}) = \mathbb{F}(\alpha)^{-1} \times \int_0^{\mathfrak{g}} t^{\alpha-1} e^{-t} dt$ for $\mathfrak{g} \geq 0$ with $\mathbb{F}(\alpha) = \int_0^{+\infty} t^{\alpha-1} e^{-t} dt$. For all $\alpha > 3$ and for all $b \in \mathbb{R}$, $h(b; \alpha) \leq 2\alpha + b^2$.

(ii) If \mathcal{N} is Gaussian with $E\{\mathcal{N}\} = 0$ and $E\{\mathcal{N}^2\} = 1$, then $E\{h(\mathcal{N}; \alpha)\} = \alpha$.

(iii) If \mathcal{G} is non-Gaussian with $E\{\mathcal{G}\} = 0$ and $E\{\mathcal{G}^2\} = 1$, we have $E\{h(\mathcal{G}; \alpha)\} \neq \alpha$, but for $\alpha \rightarrow +\infty$, $E\{h(\mathcal{G}; \alpha)\} \rightarrow \alpha$.

Proof. (Lemma 2.3). (i) Function h defined in Proposition 2.1 shows that $F_\alpha(h(b; \alpha)) = F(b)$ in which $F(b)$ is the c.d.f of \mathcal{N} and $F_\alpha(\mathfrak{g})$ is the c.d.f of the \mathbb{F}_α random variable such that $E\{\mathbb{F}_\alpha\} = \alpha$. We then deduce that $h(b; \alpha) = F_\alpha^{-1}(F(b))$. In order to prove that $h(b; \alpha) \leq 2\alpha + b^2$, since $u \mapsto F_\alpha^{-1}(u)$ is a strictly increasing function from $[0, 1[$ into \mathbb{R}^+ , we have to prove that, for all $\alpha > 3$ and for all $b \in \mathbb{R}$, we have $J_\alpha(b) \geq 0$ in which

$$(2.8) \quad J_\alpha(b) = F_\alpha(2\alpha + b^2) - F(b) = \frac{1}{\mathbb{F}(\alpha)} \int_0^{2\alpha+b^2} t^{\alpha-1} e^{-t} dt - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^b e^{-t^2/2} dt,$$

which can be rewritten as

$$(2.9) \quad J_\alpha(b) = P_\alpha + \frac{1}{\mathbb{F}(\alpha)} \int_{2\alpha}^{2\alpha+b^2} t^{\alpha-1} e^{-t} dt - \frac{\text{sgn}(b)}{\sqrt{2\pi}} \int_0^{|b|} e^{-t^2/2} dt,$$

in which $P_\alpha = F_\alpha(2\alpha) - 1/2$ that is such that, $\forall \alpha > 3$, $P_\alpha > F_3(6) - 1/2 \simeq 0.438$, and sgn is the sign function. Equation (2.9) shows that $\forall \alpha > 3$ and $\forall b < 0$, $J_\alpha(b) > 0$ because $J_\alpha(b)$ is the sum of three positive terms. For $b = 0$ and $\forall \alpha > 3$, Eq. (2.9) shows that $J_\alpha(0) = P_\alpha > 0$. For $b > 0$, we use Eq. (2.8). Let $J'_\alpha(b) = dJ_\alpha(b)/db$ be such that $J'_\alpha(b) = (2\pi)^{-1/2} e^{-b^2/2} (a_\alpha b (1 + b^2/(2\alpha))^{\alpha-1} e^{-b^2/2} - 1)$, in which $a_\alpha =$

$2\sqrt{2\pi}e^{-2\alpha}(2\alpha)^{\alpha-1}/\Gamma(\alpha)$ that is such that, $\forall \alpha > 3$, $a_\alpha \leq 0.223679$. From Eq. (2.8), it can be seen that, $\forall \alpha > 3$, for $b \rightarrow +\infty$, $J_\alpha(b) \rightarrow 0$ and $J'_\alpha(b) \sim -e^{-b^2/2}/\sqrt{2\pi}$. Consequently, $J_\alpha \rightarrow 0_+$. Since $J_\alpha(0) > 0$, we will have $J_\alpha(b) > 0$ for $b > 0$ if $J'_\alpha(b) < 0$, that is to say if $a_\alpha b(1 + \frac{b^2}{2\alpha})^{\alpha-1} < e^{b^2/2}$, which is true for $\alpha > 3$ and $b > 0$. (ii) If \mathcal{N} is normalized and Gaussian, then $h(\mathcal{N}; \alpha) = \Gamma_\alpha$ and hence $E\{h(\mathcal{N}; \alpha)\} = \alpha$. (iii) Let $P(x^2|\hat{\alpha})$ be defined by $P(x^2|\hat{\alpha}) = (2^{\hat{\alpha}/2}\Gamma(\hat{\alpha}/2))^{-1} \int_0^{x^2} v^{(\hat{\alpha}/2)-1} e^{-v/2} dv$. The change of variable $v = 2t$ yields $P(x^2|\hat{\alpha}) = (\Gamma(\hat{\alpha}/2))^{-1} \int_0^{x^2/2} t^{(\hat{\alpha}/2)-1} e^{-t} dt$. Taking $\hat{\alpha} = 2\alpha$ yields $P(x^2|\hat{\alpha}) = F_\alpha(x^2/2)$. For $\hat{\alpha} \rightarrow +\infty$, $P(x^2|\hat{\alpha}) \sim F(y)$ with $y = (x^2 - \hat{\alpha})/\sqrt{2\hat{\alpha}}$ and hence $F_\alpha(x^2/2) \sim F(y)$. Since $F_\alpha(h(b; \alpha)) = F(b)$, taking $x^2/2 = h(b; \alpha)$ yields $F(b) \sim F(y)$ that is to say $b \sim y = (2h(b; \alpha) - 2\alpha)/\sqrt{4\hat{\alpha}}$, which shows that $h(b; \alpha) \sim \alpha + b\sqrt{\alpha}$. If \mathcal{G} is a non-Gaussian random variable such that $E\{\mathcal{G}\} = 0$, then $E\{h(\mathcal{G}; \alpha)\} \sim \alpha + \sqrt{\alpha} E\{\mathcal{G}\} = \alpha$. \square

Comments on the proposed construction. The use of a nonlinear transformation of a Gaussian random vector to a non-Gaussian one in finite or infinite dimension has widely been used in the literature (see for instance [4, 7]). The transformation presented in Proposition 2.1 is not trivial because it does not result from a simple nonlinear transformation of a Gaussian vector but takes into account the matrix structure and its algebraic properties. This nonlinear transformation, which has been constructed by using the Maximum Entropy principle, appears as the composition of two nonlinear transformations for which, the second one (transformation h) depends on parameters linked to the components of the random matrix. This construction was introduced in [28, 29], but in the present paper, point (i) of Lemma 2.3 is a new result, which is essential to demonstrate the existence and uniqueness of the strong stochastic solution of the weak formulation of the stochastic elliptic boundary value problem for stochastic homogenization (see Sections 4 and 5).

Remark 2.4. The analysis of the proof of Proposition 2.1 shows that the property $E\{[\mathbf{C}(\mathbf{x}; \mathbf{S})]\} = [I_6]$ holds for a fixed value of \mathbf{S} . When \mathbf{S} will be modeled by a random variable \mathbf{S} in order to take into account uncertainties in the spectral measure (see Section 4 and as we have previously explained at the beginning of Section 2) the random field $\{[\tilde{\mathbf{C}}(\mathbf{x})], \mathbf{x} \in \mathbb{R}^3\}$ such that $[\tilde{\mathbf{C}}(\mathbf{x})] = [\mathbf{C}(\mathbf{x}; \mathbf{S})]$ will then depend on copies of the random field $\{G(\mathbf{x}; \mathbf{S}), \mathbf{x} \in \mathbb{R}^3\}$ that will always satisfy $E\{G(\mathbf{x}; \mathbf{S})\} = 0$ and $E\{G(\mathbf{x}; \mathbf{S})^2\} = 1$, but which will no longer be Gaussian. By examining the proof of Proposition 2.1, it can be seen that we will always have $E\{[\tilde{\mathbf{C}}(\mathbf{x})]_{mn}\} = 0$ for $m \neq n$ but that $E\{[\tilde{\mathbf{C}}(\mathbf{x})]_{mm}\} = 2\sigma_c^2 E\{h(G(\mathbf{x}; \mathbf{S}); \alpha_m)\} + \sigma_c^2(m-1) \neq 1$. Nevertheless, from Proposition 2.1, since $\alpha_m > 3$ and from Lemma 2.3-(iii), taking $\mathcal{G} = G(\mathbf{x}; \mathbf{S})$ yields $E\{[\tilde{\mathbf{C}}(\mathbf{x})]_{mm}\} \simeq 1$. It has numerically been verified that, if \mathcal{G} is a uniform random variable (that will not be the case for $G(\mathbf{x}; \mathbf{S})$) such that $E\{\mathcal{G}\} = 0$ and $E\{\mathcal{G}^2\} = 1$ (that will be the case for $G(\mathbf{x}; \mathbf{S})$), then we have $E\{h(\mathcal{G}; \alpha_m)\} \simeq \alpha_m$ with an error of 5×10^{-4} for all $1 \leq m \leq 6$.

3. CONSTRUCTION AND ANALYSIS OF THE GAUSSIAN RANDOM FIELD $G(\cdot; \mathbf{S})$ WITH UNCERTAIN SPECTRAL MEASURE PARAMETERIZED BY \mathbf{S}

We start by constructing a normalized Gaussian, homogeneous, second-order, mean-square continuous random field $\{G(\mathbf{x}), \mathbf{x} \in \mathbb{R}^3\}$. This field corresponds to $G(\cdot; \underline{\mathbf{S}})$ for which its spectral measure is given and represented by a given value $\underline{\mathbf{S}}$ of \mathbf{S} , that is to say, $G = G(\cdot; \underline{\mathbf{S}})$. Therefore, there exists a positive bounded spectral measure $m_G(d\mathbf{k})$ on \mathbb{R}^3 such that the correlation function ρ_G of G is written, for all \mathbf{x} and $\boldsymbol{\zeta}$ in \mathbb{R}^3 , as

$$(3.1) \quad \rho_G(\boldsymbol{\zeta}) = E\{G(\mathbf{x} + \boldsymbol{\zeta})G(\mathbf{x})\} = \int_{\mathbb{R}^3} e^{i\mathbf{k} \cdot \boldsymbol{\zeta}} m_G(d\mathbf{k}) = \int_{\mathbb{R}^3} \cos(\mathbf{k} \cdot \boldsymbol{\zeta}) m_G(d\mathbf{k}),$$

in which $\iota = \sqrt{-1}$, $\mathbf{k} \cdot \boldsymbol{\zeta} = \sum_{j=1}^3 k_j \zeta_j$, $d\mathbf{k} = dk_1 dk_2 dk_3$. In addition, it is assumed that $m_G(d\mathbf{k}) = s(\mathbf{k}) d\mathbf{k}$ admits a spectral density function (s.d.f) $\mathbf{k} \mapsto s(\mathbf{k}) : \mathbb{R}^3 \rightarrow \mathbb{R}^+$. Equations (2.7) and (3.1) yield $\rho_G(0) = E\{G(\mathbf{x})^2\} = 1$, and consequently,

$$(3.2) \quad \int_{\mathbb{R}^3} s(\mathbf{k}) d\mathbf{k} = 1.$$

In this section, we begin with the analysis of the spectral measure and the modeling of uncertainties. Then we introduce a finite representation of $G(\cdot; \mathcal{S})$ and we study its properties. Note that a dimensionless s.d.f, χ , of s will be introduced.

Hypothesis 2 (*Spectral density function s and spatial correlation length of G*). It is assumed that s has a compact support $\bar{\mathbb{K}} = \partial\mathbb{K} \cup \mathbb{K}$ with $\mathbb{K} = \prod_{j=1}^3]-K_j, K_j[$ in which $K_j \in [K_j^{\min}, K_j^{\max}]$ with $0 < K_j^{\min} < K_j^{\max} < +\infty$. It is assumed that s is a continuous function on \mathbb{R}^3 . Since $\text{supp } s = \bar{\mathbb{K}}$, we must have $s(\mathbf{k}) = 0, \forall \mathbf{k} \in \partial\mathbb{K}$, and thus

$$(3.3) \quad \int_{\mathbb{R}^3} s(\mathbf{k}) d\mathbf{k} = \int_{\bar{\mathbb{K}}} s(\mathbf{k}) d\mathbf{k} = 1.$$

Since G is real, we have $s(-\mathbf{k}) = s(\mathbf{k})$ for all $\mathbf{k} \in \mathbb{R}^3$. In addition to this symmetry property, we assume that s satisfies the following quadrant symmetry [34]: defining $\mathbf{k}_{\{-j\}}$ as vector \mathbf{k} for which its component k_j is replaced by $-k_j$, then $s(\mathbf{k}_{\{-j\}}) = s(\mathbf{k})$ for $j = 1, 2, 3$ and $\forall \mathbf{k} \in \mathbb{R}^3$. The spatial correlation length for coordinate ζ_j is defined by

$$(3.4) \quad L_{cj} = \int_0^{+\infty} |\rho_j(0, \dots, \zeta_j, \dots, 0)| d\zeta_j,$$

and is assumed to be finite.

Definition 3.1 (*Spectral domain sampling*). Let ν_s be a given even integer. For $j \in \{1, 2, 3\}$, we define $\Delta_j = 2K_j/\nu_s$ as the sampling step of interval $[-K_j, K_j]$ and $k_{j\beta_j} = -K_j + (\beta_j - 1/2)\Delta_j$ for $\beta_j = 1, \dots, \nu_s$ as its spectral sampling points. Let $\mathcal{B} = \{\boldsymbol{\beta} = (\beta_1, \beta_2, \beta_3), \beta_j = 1, \dots, \nu_s\}$ be the finite subset of \mathbb{N}^3 . We define Δ, K, ν , and \mathbf{k}_β such that: $\Delta = \Delta_1 \Delta_2 \Delta_3, K = K_1 K_2 K_3, \nu = (\nu_s)^3$, and for all $\beta \in \mathcal{B}, \mathbf{k}_\beta = (k_{1\beta_1}, k_{2\beta_2}, k_{3\beta_3}) \in \mathbb{K} \subset \mathbb{R}^3$.

Lemma 3.2 (*Discretization of the spectral measure and convergence properties*). Let $\delta_{\mathbf{k}_\beta}(\mathbf{k}) = \otimes_{j=1}^3 \delta_{k_{j\beta_j}}(k_j)$ be the Dirac measure on \mathbb{R}^3 at sampling point $\mathbf{k}_\beta \in \mathbb{K} \subset \mathbb{R}^3$ defined in Definition 3.1. Let $m_G^\nu(d\mathbf{k})$ be the positive bounded measure on \mathbb{R}^3 defined by

$$(3.5) \quad m_G^\nu(d\mathbf{k}) = \sum_{\beta \in \mathcal{B}} s_\beta^\Delta \delta_{\mathbf{k}_\beta}(\mathbf{k}) \quad , \quad s_\beta^\Delta = \Delta s(\mathbf{k}_\beta),$$

which is such that $m_G^\nu(\mathbb{R}^3) = \sum_{\beta \in \mathcal{B}} s_\beta^\Delta = \eta_\nu$ with $\eta_\nu > 0$. The sequence of measures $\{m_G^\nu(d\mathbf{k})\}_\nu$ converges narrowly towards the measure $m_G(d\mathbf{k})$ and the positive sequence $\{\eta_\nu\}_\nu$ converges towards 1.

Proof. (Lemma 3.2). We have to prove that $\forall f \in C^0(\bar{\mathbb{K}})$, the sequence $m_G^\nu(f) = \int_{\bar{\mathbb{K}}} f(\mathbf{k}) m_G^\nu(d\mathbf{k}) = \Delta \sum_{\beta \in \mathcal{B}} f(\mathbf{k}_\beta) s(\mathbf{k}_\beta)$ converges towards $m_G(f) = \int_{\bar{\mathbb{K}}} f(\mathbf{k}) m_G(d\mathbf{k}) = \int_{\bar{\mathbb{K}}} f(\mathbf{k}) s(\mathbf{k}) d\mathbf{k}$. Since the function $\mathbf{k} \mapsto f(\mathbf{k}) s(\mathbf{k})$ is continuous on $\bar{\mathbb{K}}$, it is known that for $\nu_s \rightarrow +\infty$ (that is to say for $\nu \rightarrow +\infty$), $\Delta \sum_{\beta \in \mathcal{B}} f(\mathbf{k}_\beta) s(\mathbf{k}_\beta) \rightarrow \int_{\bar{\mathbb{K}}} f(\mathbf{k}) s(\mathbf{k}) d\mathbf{k}$. Taking $f(\mathbf{k}) = 1$ for all $\mathbf{k} \in \bar{\mathbb{K}}$ yields $m_G(1) = \int_{\bar{\mathbb{K}}} s(\mathbf{k}) d\mathbf{k} = 1$ and $m_G^\nu(1) = \Delta \sum_{\beta \in \mathcal{B}} s(\mathbf{k}_\beta) = \sum_{\beta \in \mathcal{B}} s_\beta^\Delta = \eta_\nu$. Therefore, $\{\eta_\nu\}_\nu$ converges towards 1. \square

Hypothesis 3 (*Choice of $\nu = (\nu_s)^3$*). Let s be the s.d.f satisfying Hypothesis 2. Let us consider the spectral domain sampling introduced in Definition 3.1. Using Lemma 3.2,

we will assume that ν is chosen sufficiently large in order that $|\sum_{\beta \in \mathcal{B}} s_{\beta}^{\Delta} - 1| \leq \epsilon_s \ll 1$ and consequently, we will write $\sum_{\beta \in \mathcal{B}} s_{\beta}^{\Delta} \simeq 1$.

Definition 3.3 (*Dimensionless spectral density function*). The spectral density function $\mathbf{k} \mapsto s(\mathbf{k})$, which verifies Hypothesis 2, is written for all \mathbf{k} in \mathbb{R}^3 as $s(k_1, k_2, k_3) = (K_1 K_2 K_3)^{-1} \chi(k_1/K_1, k_2/K_2, k_3/K_3)$ in which χ is a given function $\boldsymbol{\tau} = (\tau_1, \tau_2, \tau_3) \mapsto \chi(\tau_1, \tau_2, \tau_3) : \mathbb{R}^3 \rightarrow \mathbb{R}^+$ with compact support $[-1, 1]^3$.

Function χ has the same properties as s : $\chi(-\boldsymbol{\tau}) = \chi(\boldsymbol{\tau})$, quadrant symmetry, and continuity. For $j = 1, 2, 3$, the change of variable $\tau_j = k_j/K_j$ yields $s(\mathbf{k}) d\mathbf{k} = \chi(\boldsymbol{\tau}) d\boldsymbol{\tau}$ and thus $m_G(d\mathbf{k}) = \mu_G(d\boldsymbol{\tau})$ with $\mu_G(d\boldsymbol{\tau}) = \chi(\boldsymbol{\tau}) d\boldsymbol{\tau}$, and consequently, Eq. (3.3) yields

$$(3.6) \quad \int_{\mathbb{R}^3} \chi(\boldsymbol{\tau}) d\boldsymbol{\tau} = \int_{[-1, 1]^3} \chi(\boldsymbol{\tau}) d\boldsymbol{\tau} = 1.$$

The dimensionless spectral domain sampling is directly deduced from Definition 3.1,

$$(3.7) \quad \{\boldsymbol{\tau}_{\beta} = (\tau_{\beta_1}, \tau_{\beta_2}, \tau_{\beta_3}), \boldsymbol{\beta} \in \mathcal{B}\}, \tau_{\beta_j} = -1 + (\beta_j - \frac{1}{2}) \frac{2}{\nu_s}, j \in \{1, 2, 3\}.$$

The discretization $\mu_G^{\nu}(d\boldsymbol{\tau})$ of $\mu_G(d\boldsymbol{\tau})$, such that $\mu_G^{\nu}(d\boldsymbol{\tau}) = m_G^{\nu}(d\mathbf{k})$, is written as

$$(3.8) \quad \mu_G^{\nu}(d\boldsymbol{\tau}) = \sum_{\boldsymbol{\beta} \in \mathcal{B}} \chi_{\boldsymbol{\beta}}^{\Delta} \delta_{\boldsymbol{\tau}_{\boldsymbol{\beta}}}(\boldsymbol{\tau}) \quad , \quad \chi_{\boldsymbol{\beta}}^{\Delta} = (2/\nu_s)^3 \chi(\boldsymbol{\tau}_{\boldsymbol{\beta}}),$$

in which $\delta_{\boldsymbol{\tau}_{\boldsymbol{\beta}}} = \otimes_{j=1}^3 \delta_{\tau_{\beta_j}}(\tau_j)$ and where, from Hypothesis 3,

$$(3.9) \quad \sum_{\boldsymbol{\beta} \in \mathcal{B}} \chi_{\boldsymbol{\beta}}^{\Delta} \simeq 1.$$

Definition 3.3 implies that measure $\mu_G^{\nu}(d\boldsymbol{\tau})$ is independent of K_1, K_2 , and K_3 . In order to introduce the probability model of the spectral measure, we start by defining an adapted parameterization $[y]$ that takes into account quadrant symmetry.

Definition 3.4 (*Parameterization of the discretized dimensionless spectral measure*). Let $\widehat{\nu}_s = \nu_s/2$ (ν_s is even). Let \mathcal{C}_y be the subset of $\mathcal{M}_{3, \widehat{\nu}_s}$ defined by

$$(3.10) \quad \mathcal{C}_y = \{ [y] \in \mathcal{M}_{3, \widehat{\nu}_s}, [y]_{j\widehat{\beta}} \in [0, 1] \text{ for } j = 1, 2, 3 \text{ and } \widehat{\beta} = 1, \dots, \widehat{\nu}_s \}.$$

Let $[y]$ be in \mathcal{C}_y such that $[y]_{j\widehat{\beta}} = 1/2$ for $j = 1, 2, 3$ and $\widehat{\beta} = 1, \dots, \widehat{\nu}_s$.

Let $\widehat{\mathcal{B}} = \{\widehat{\boldsymbol{\beta}} = (\widehat{\beta}_1, \widehat{\beta}_2, \widehat{\beta}_3), \widehat{\beta}_j = 1, \dots, \widehat{\nu}_s\} \subset \mathcal{B}$ be the set of $\widehat{\nu} = (\widehat{\nu}_s)^3 = \nu/8$ elements. We define the finite family of functions $[y] \mapsto a_{\widehat{\boldsymbol{\beta}}}([y]) : \mathcal{C}_y \rightarrow \mathbb{R}$ such that $a_{\widehat{\boldsymbol{\beta}}}([y]) = \sqrt{\chi_{\widehat{\boldsymbol{\beta}}}^{\Delta} q_{\widehat{\boldsymbol{\beta}}}([y]; \delta_s)}$, in which $\delta_s > 0$ is a hyperparameter that will allow the level of spectrum uncertainties to be controlled and where $[y] \mapsto q_{\widehat{\boldsymbol{\beta}}}([y]; \delta_s)$ is any given continuous real function on \mathcal{C}_y such that $q_{\widehat{\boldsymbol{\beta}}}([y]; \delta_s) = 1$. For all $[y] \in \mathcal{C}_y$, let $\{a_{\boldsymbol{\beta}}([y]), \boldsymbol{\beta} \in \mathcal{B}\}$ be the ν real numbers that are directly constructed from $\{a_{\widehat{\boldsymbol{\beta}}}([y]), \widehat{\boldsymbol{\beta}} \in \widehat{\mathcal{B}}\}$ using the quadrant symmetry (see Hypothesis 2); an example of such a construction is given in Example 3.5-(iv). For all $\boldsymbol{\beta} \in \mathcal{B}$, we define the function $[y] \mapsto \widetilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([y]) : \mathcal{C}_y \rightarrow \mathbb{R}^+$ such that

$$(3.11) \quad \widetilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([y]) = a_{\boldsymbol{\beta}}([y])^2 \left(\sum_{\boldsymbol{\beta}' \in \mathcal{B}} a_{\boldsymbol{\beta}'}([y])^2 \right)^{-1}.$$

The dimensionless spectral measure $\widetilde{\mu}_G^{\nu}(d\boldsymbol{\tau}; [y])$ for $[y]$ given in \mathcal{C}_y is then defined by

$$(3.12) \quad \widetilde{\mu}_G^{\nu}(d\boldsymbol{\tau}; [y]) = \sum_{\boldsymbol{\beta} \in \mathcal{B}} \widetilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([y]) \delta_{\boldsymbol{\tau}_{\boldsymbol{\beta}}}(\boldsymbol{\tau}).$$

Comments concerning the parameterization of uncertainties in the spectral measure. The spectral density function $\mathbf{k} \mapsto s(\mathbf{k})$ has a compact support $\bar{\mathbb{K}} = \prod_{j=1}^3 [-K_j, K_j] \subset \mathbb{R}^3$ and has been written as $s(k_1, k_2, k_3) = (K_1 K_2 K_3)^{-1} \chi(k_1/K_1, k_2/K_2, k_3/K_3)$ in which χ is the dimensionless spectral density function, $\boldsymbol{\tau} = (\tau_1, \tau_2, \tau_3) \mapsto \chi(\tau_1, \tau_2, \tau_3)$, with compact support $[-1, 1]^3$. The considered uncertainties are one hand on the support $\bar{\mathbb{K}}$ and on the other hand on the shape of dimensionless spectral density function χ on its support, which is introduced by a family $\{\chi(\cdot; [y]), [y] \in \mathcal{C}_y\}$ of dimensionless spectral density functions. Introducing $\mathbf{w} = (w_1, w_2, w_3) \in \mathcal{C}_w \subset \mathbb{R}^3$ with $w_j = \pi/K_j$ for $j = 1, 2, 3$, the parameter $\mathbf{S} = \{\mathbf{w}, [y]\} \in \mathcal{C}_S = \mathcal{C}_w \times \mathcal{C}_y$ is modeled by the random variable $\mathbf{S} = \{\mathbf{W}, [\mathbf{Y}]\}$ whose support of its probability measure is $\mathcal{C}_S = \mathcal{C}_w \times \mathcal{C}_y$. The mean value $\underline{\mathbf{S}}$ of \mathbf{S} is $\{\underline{\mathbf{w}}, [y]\}$.

Proposition 3.1 (Random discretized dimensionless spectral measure). *Let us consider Definitions 3.3 and 3.4. It is assumed that Eq. (3.9) holds.*

(i) $\forall \boldsymbol{\beta} \in \mathcal{B}$, function $[y] \mapsto \tilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([y])$ is continuous on \mathcal{C}_y (and thus bounded on \mathcal{C}_y), is such that $\tilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([y]) \simeq \chi_{\boldsymbol{\beta}}^{\Delta}$, and $\forall [y] \in \mathcal{C}_y$, $\sum_{\boldsymbol{\beta} \in \mathcal{B}} \tilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([y]) = 1$.

(ii) Let $[\mathbf{Y}]$ be the $\mathbb{M}_{3, \hat{\nu}_s}$ -valued random variable, defined on $(\Theta, \mathcal{T}, \mathcal{P})$, whose support of its probability measure is $\mathcal{C}_y \subset \mathbb{M}_{3, \hat{\nu}_s}$, and such that $\{[\mathbf{Y}]_{j\hat{\beta}}, j \in \{1, 2, 3\}, \hat{\beta} \in \{1, \dots, \hat{\nu}_s\}\}$ are $3\hat{\nu}_s$ independent uniform random variables on $[0, 1]$. Its mean value is $E\{[\mathbf{Y}]\} = \int_{\mathcal{C}_y} [y] P_{[\mathbf{Y}]}(dy) = \int_{\mathcal{C}_y} [y] dy = [y]$. For all $\hat{\boldsymbol{\beta}} \in \hat{\mathcal{B}}$, $A_{\hat{\boldsymbol{\beta}}} = a_{\hat{\boldsymbol{\beta}}}([\mathbf{Y}])$ is a second-order real-valued random variable.

(iii) $\forall \boldsymbol{\beta} \in \mathcal{B}$, $\tilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([\mathbf{Y}])$ is a second-order positive-valued random variable, defined on $(\Theta, \mathcal{T}, \mathcal{P})$ such that $\sum_{\boldsymbol{\beta} \in \mathcal{B}} \tilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([\mathbf{Y}]) = 1$ almost surely.

(iv) The dimensionless spectral measure $\tilde{\mu}_G^{\nu}(\mathbf{d}\boldsymbol{\tau}; [y])$ for given $[y]$ in \mathcal{C}_y , is a bounded positive measure on \mathbb{R}^3 and is such that $\tilde{\mu}_G^{\nu}(\mathbf{d}\boldsymbol{\tau}; [y]) \simeq \mu_G^{\nu}(\mathbf{d}\boldsymbol{\tau})$. For all $[y] \in \mathcal{C}_y$, $\tilde{\mu}_G^{\nu}(\mathbb{R}^3; [y]) = \sum_{\boldsymbol{\beta} \in \mathcal{B}} \tilde{\chi}_{\boldsymbol{\beta}}^{\Delta}([y]) = 1$.

Proof. (Proposition 3.1). This proposition is easy to prove and is left to the reader. \square

Example 3.5 (Illustration of a construction for a separable spatial correlation structure).

(i) *Spectral density function.* $\forall \mathbf{k} = (k_1, k_2, k_3) \in \mathbb{R}^3$, $s(\mathbf{k}) = \prod_{j=1}^3 s_j(k_j)$. For $j = 1, 2, 3$, $s_j(k_j) = K_j^{-1} (1 - |k_j|/K_j) \mathbb{1}_{[-K_j, K_j]}(k_j)$ and thus, $\text{supp } s_j = [-K_j, K_j]$, $s_j(-k_j) = s_j(k_j)$ (yielding $s(-\mathbf{k}) = s(\mathbf{k})$ and the quadrant symmetry), and $\int_{[-K_j, K_j]} s_j(k_j) dk_j = 1$.

(ii) *Correlation function and spatial correlation length.* For all $\boldsymbol{\zeta} = (\zeta_1, \zeta_2, \zeta_3) \in \mathbb{R}^3$, $\rho_G(\boldsymbol{\zeta}) = \prod_{j=1}^3 \rho_j(\zeta_j)$ and for $j = 1, 2, 3$, $\rho_j(\zeta_j) = \int_{\mathbb{R}} e^{i k_j \zeta_j} s_j(k_j) dk_j$, $\rho_j(0) = 1$, and the spatial correlation length is $L_{c_j} = \int_0^{+\infty} |\rho_j(\zeta_j)| d\zeta_j = \pi s_j(0) = \pi/K_j$.

(iii) *Dimensionless spectral density function and spectral sampling.* $\forall \boldsymbol{\tau} = (\tau_1, \tau_2, \tau_3)$, $\chi(\boldsymbol{\tau}) = \prod_{j=1}^3 \chi_j(\tau_j)$. For $j = 1, 2, 3$, $\chi_j(\tau_j) = (1 - |\tau_j|) \mathbb{1}_{[-1, 1]}(\tau_j)$ and therefore, $\text{supp } \chi_j = [-1, 1]$, $\chi_j(-\tau_j) = \chi_j(\tau_j)$, and $\int_{\mathbb{R}} \chi_j(\tau_j) d\tau_j = 1$. For all $\boldsymbol{\beta} = (\beta_1, \beta_2, \beta_3) \in \mathcal{B}$, $\chi_{\boldsymbol{\beta}}^{\Delta} = \prod_{j=1}^3 \chi_{j\beta_j}^{\Delta}$ with $\chi_{j\beta_j}^{\Delta} = (2/\nu_s) \chi_j(\tau_{\beta_j})$.

(iv) *Construction of $a_{\boldsymbol{\beta}}([y])$.* For $j \in \{1, 2, 3\}$, $\hat{\beta}_j \in \{1, \dots, \hat{\nu}_s\}$, $\hat{\boldsymbol{\beta}} = (\hat{\beta}_1, \hat{\beta}_2, \hat{\beta}_3)$, and $\forall [y] \in \mathcal{C}_y$, $q_{\hat{\boldsymbol{\beta}}}([y]; \delta_s) = \prod_{j=1}^3 q_{j\hat{\beta}_j}([y]; \delta_j)$ in which $q_{j\hat{\beta}_j}([y]; \delta_j) = 1 + \sqrt{12} \delta_j ([y]_{j\hat{\beta}_j} - 1/2)$ with $\delta_j > 0$ the hyperparameter. We thus have $a_{\hat{\boldsymbol{\beta}}}([y]) = \prod_{j=1}^3 a_{j\hat{\beta}_j}([y])$ in which $a_{j\hat{\beta}_j}([y]) = \sqrt{\chi_{j\hat{\beta}_j}^{\Delta}} q_{j\hat{\beta}_j}([y]; \delta_j)$ and for $\beta_j \in \{\hat{\nu}_s+1, \dots, 2\hat{\nu}_s\}$, $a_{j\beta_j}([y]) = a_{j(2\hat{\nu}_s+1-\beta_j)}([y])$.

(v) *Random variable $A_{\hat{\boldsymbol{\beta}}}$ and hyperparameter δ_s .* For $j \in \{1, 2, 3\}$ and $\hat{\beta}_j \in \{1, \dots, \hat{\nu}_s\}$, the mean value and the second-order moment of random variable $A_{j\hat{\beta}_j} = a_{j\hat{\beta}_j}([\mathbf{Y}])$ are $E\{A_{j\hat{\beta}_j}\} = \sqrt{\chi_{j\hat{\beta}_j}^{\Delta}}$ and $E\{A_{j\hat{\beta}_j}^2\} = \chi_{j\hat{\beta}_j}^{\Delta} (1 + \delta_j^2)$. Since the random variables $\{A_{j\hat{\beta}_j}\}_{j, \hat{\beta}_j}$ are independent, the mean value and the second-order moment of the random variable

$A_{\hat{\beta}} = a_{\hat{\beta}}([\mathbf{Y}]) = \prod_{j=1}^3 A_{j\hat{\beta}_j}$ are $E\{A_{\hat{\beta}}\} = \sqrt{\chi_{\hat{\beta}}^{\Delta}}$ and $E\{A_{\hat{\beta}}^2\} = \chi_{\hat{\beta}}^{\Delta} \prod_{j=1}^3 (1 + \delta_j^2)$. Defining the hyperparameter δ_s as $\delta_s^2 = E\{(A_{\hat{\beta}} - \sqrt{\chi_{\hat{\beta}}^{\Delta}})^2\} / \chi_{\hat{\beta}}^{\Delta}$, it can be seen that we have $\delta_s^2 = (\prod_{j=1}^3 (1 + \delta_j^2)) - 1 > 0$, which is independent of $\hat{\beta}$.

(vi) *Discretized dimensionless spectral measure.* Eq. (3.11) yields, $\forall \beta = (\beta_1, \beta_2, \beta_3) \in \mathcal{B}$, $\tilde{\chi}_{\beta}^{\Delta}([y]) = \prod_{j=1}^3 \tilde{\chi}_{j\beta_j}^{\Delta}([y])$ in which $\tilde{\chi}_{j\beta_j}^{\Delta}([y]) = a_{j\beta_j}([y])^2 (\sum_{\beta' \in \mathcal{B}} a_{\beta'}([y])^2)^{-1/3}$.

Definition 3.6 (*Spectrum parameters \mathbf{w} and \mathcal{S} and their probabilistic models \mathbf{W} and \mathcal{S}*). Let $\mathbf{w} = (w_1, w_2, w_3)$ in which $w_j = \pi/K_j$ (this parameter allows the support of the spectral measure to be controlled). Let $\mathcal{C}_w = \{\mathbf{w} \in \mathbb{R}^3; w_j \in [w_j^{\min}, w_j^{\max}]\}$ for $j = 1, 2, 3$ be the compact subset of \mathbb{R}^3 , in which $0 < w_j^{\min} = \pi/K_j^{\max} < w_j^{\max} = \pi/K_j^{\min} < +\infty$ (see Hypothesis 2). Parameter \mathbf{w} is modeled by a \mathbb{R}^3 -valued random variable \mathbf{W} , defined on $(\Theta, \mathcal{T}, \mathcal{P})$, independent of $[\mathbf{Y}]$, whose support of its given probability measure $P_{\mathbf{W}}(d\mathbf{w})$ is \mathcal{C}_w . We define the parameter \mathcal{S} as $\{\mathbf{w}, [y]\}$, which takes its values in the subset $\mathcal{C}_{\mathcal{S}} = \mathcal{C}_w \times \mathcal{C}_y$ of $\mathbb{R}^3 \times \mathbb{M}_{3, \mathcal{D}_s}$. The probabilistic model of \mathcal{S} is the $\mathbb{R}^3 \times \mathbb{M}_{3, \mathcal{D}_s}$ -valued random variable $\mathcal{S} = \{\mathbf{W}, [\mathbf{Y}]\}$ whose probability measure is the product of measures $P_{\mathcal{S}} = P_{\mathbf{W}}(d\mathbf{w}) \otimes P_{[\mathbf{Y}]}(dy)$ whose compact support is $\mathcal{C}_{\mathcal{S}}$.

Definition 3.7 (*Normalized Gaussian random field $G^{\nu}(\cdot; \mathcal{S})$ given $\mathcal{S} = \{\mathbf{w}, [y]\} \in \mathcal{C}_{\mathcal{S}} = \mathcal{C}_w \times \mathcal{C}_y$*). Let $\nu = (\nu_s)^3$ be fixed. Let $\{Z_{\beta}, \beta \in \mathcal{B}\}$ and $\{\Phi_{\beta}, \beta \in \mathcal{B}\}$ be 2ν independent random variables on $(\Theta, \mathcal{T}, \mathcal{P})$, which are independent of \mathbf{W} and $[\mathbf{Y}]$. For all $\beta \in \mathcal{B}$, $Z_{\beta} = \sqrt{-\log \Psi_{\beta}}$ in which Ψ_{β} is uniform of $[0, 1]$ and Φ_{β} is uniform on $[0, 2\pi]$. Let $P_{\mathbf{Z}}(dz)$ and $P_{\Phi}(d\varphi)$ be the probability measures on \mathbb{R}^{ν} of the \mathbb{R}^{ν} -valued random variables $\mathbf{Z} = \{Z_{\beta}, \beta \in \mathcal{B}\}$ and $\Phi = \{\Phi_{\beta}, \beta \in \mathcal{B}\}$. The unbounded support \mathcal{C}_z of $P_{\mathbf{Z}}(dz)$ is $\mathcal{C}_z = \{z = \{z_{\beta}, \beta \in \mathcal{B}\}, z_{\beta} > 0\} \subset \mathbb{R}^{\nu}$ and the compact support \mathcal{C}_{φ} of $P_{\Phi}(d\varphi)$ is $\mathcal{C}_{\varphi} = \{\varphi = \{\varphi_{\beta}, \beta \in \mathcal{B}\}, \varphi_{\beta} \in [0, 2\pi]\} \subset \mathbb{R}^{\nu}$. Let $\mathbf{x} \mapsto g^{\nu}(\mathbf{x}; \mathcal{S}, \mathbf{z}, \varphi) : \mathbb{R}^3 \rightarrow \mathbb{R}$ be such that, for all $\{\mathcal{S}, \mathbf{z}, \varphi\} \in \mathcal{C}_{\mathcal{S}} \times \mathcal{C}_z \times \mathcal{C}_{\varphi}$,

$$(3.13) \quad g^{\nu}(\mathbf{x}; \mathcal{S}, \mathbf{z}, \varphi) = \sum_{\beta \in \mathcal{B}} \sqrt{2\tilde{\chi}_{\beta}^{\Delta}([y])} z_{\beta} \cos(\varphi_{\beta} + \sum_{j=1}^3 \frac{\pi}{w_j} \tau_{\beta_j} x_j).$$

For all $\mathbf{w} \in \mathcal{C}_w$ and $[y] \in \mathcal{C}_y$, we define the real-valued random field $\{G^{\nu}(\mathbf{x}; \mathcal{S}), \mathbf{x} \in \mathbb{R}^3\}$ with $\mathcal{S} = \{\mathbf{w}, [y]\} \in \mathcal{C}_{\mathcal{S}} = \mathcal{C}_w \times \mathcal{C}_y$, such that

$$(3.14) \quad G^{\nu}(\mathbf{x}; \mathcal{S}) = g^{\nu}(\mathbf{x}; \mathcal{S}, \mathbf{Z}, \Phi).$$

Equation (3.14) with Eq. (3.13) corresponds to a finite discretization of the stochastic integral representation with a stochastic spectral measure for homogeneous second-order mean-square continuous random fields [6, 8, 12].

Proposition 3.2 (*Properties of random field $G^{\nu}(\cdot; \mathcal{S})$*). For all $\mathcal{S} = \{\mathbf{w}, [y]\} \in \mathcal{C}_{\mathcal{S}} = \mathcal{C}_w \times \mathcal{C}_y$, the real-valued random field $\{G^{\nu}(\mathbf{x}; \mathcal{S}), \mathbf{x} \in \mathbb{R}^3\}$ is Gaussian, homogeneous, second-order, mean-square continuous, and normalized,

$$(3.15) \quad E\{G^{\nu}(\mathbf{x}; \mathcal{S})\} = 0 \quad , \quad E\{G^{\nu}(\mathbf{x}; \mathcal{S})^2\} = 1 \quad , \quad \forall \mathbf{x} \in \mathbb{R}^3.$$

Its dimensionless spectral measure $\tilde{\mu}_G^{\nu}(d\boldsymbol{\tau}; [y])$, expressed with the dimensionless spectral variable $\boldsymbol{\tau} = (\tau_1, \tau_2, \tau_3)$ with $\tau_j = k_j/K_j$, is the spectral measure defined by Eq. (3.12).

Proof. (Proposition 3.2). Since $\{Z_{\beta}, \beta \in \mathcal{B}\}$ and $\{\Phi_{\beta}, \beta \in \mathcal{B}\}$ are 2ν independent random variables, it can easily be proven that $G^{\nu}(\cdot; \mathcal{S})$ is centered (first equation in Eq. (3.15)) and, $\forall \boldsymbol{\zeta} = (\zeta_1, \zeta_2, \zeta_3) \in \mathbb{R}^3$ and $\forall \mathcal{S} = \{\mathbf{w}, [y]\} \in \mathcal{C}_{\mathcal{S}} = \mathcal{C}_w \times \mathcal{C}_y$,

$$(3.16) \quad \rho_G^{\nu}(\boldsymbol{\zeta}; \mathcal{S}) = E\{G^{\nu}(\mathbf{x} + \boldsymbol{\zeta}; \mathcal{S}) G^{\nu}(\mathbf{x}; \mathcal{S})\} = \sum_{\beta \in \mathcal{B}} \tilde{\chi}_{\beta}^{\Delta}([y]) \cos(\sum_{j=1}^3 \frac{\pi}{w_j} \tau_{\beta_j} \zeta_j).$$

Using Proposition 3.1-(i) yields the second equation in Eq. (3.15). For all $\beta \in \mathcal{B}$, the random variable $Z_\beta \cos(\Phi_\beta + \sum_{j=1}^3 \frac{\pi}{w_j} \tau_{\beta_j} x_j)$ is Gaussian and consequently, random field $G^\nu(\cdot; \mathcal{S})$ is Gaussian. Since $G^\nu(\cdot; \mathcal{S})$ is a Gaussian random field with zero mean function and a correlation function that depends only on ζ , $G^\nu(\cdot; \mathcal{S})$ is homogeneous on \mathbb{R}^3 . Since $\zeta \mapsto \rho_G^\nu(\zeta; \mathcal{S})$ defined by Eq. (3.16) is continuous on \mathbb{R}^3 , $G^\nu(\cdot; \mathcal{S})$ is mean-square continuous on \mathbb{R}^3 and thus there exists a spectral measure given by Eq. (3.12). Note that the spectral measure in $\mathbf{k} = (k_1, k_2, k_3)$ is such that $\tilde{m}_G^\nu(d\mathbf{k}; \mathcal{S}) = \tilde{\mu}_G^\nu(d\boldsymbol{\tau}; [y])$ with $\tilde{m}_G^\nu(d\mathbf{k}; \mathcal{S}) = \sum_{\beta \in \mathcal{B}} \tilde{s}_\beta^\Delta(\boldsymbol{\omega}, [y]) \delta_{\mathbf{k}_\beta}(\mathbf{k})$ in which $\tilde{s}_\beta^\Delta(\boldsymbol{\omega}, [y]) = \tilde{\chi}_\beta^\Delta([y]) \Pi_{j=1}^3(\pi/w_j)$ with $\mathcal{S} = \{\boldsymbol{\omega}, [y]\}$. \square

4. NON-GAUSSIAN RANDOM FIELD $[\mathbf{C}(\cdot; \mathcal{S})]$ PARAMETERIZED BY \mathcal{S} AND RANDOM FIELD $[\tilde{\mathbf{C}}]$ WITH UNCERTAIN SPECTRAL MEASURE

In the construction of G with an uncertain spectrum, the spectral measure $m_G(d\mathbf{k})$ of G is given (see Hypothesis 2). This is the reason why the convergence of the sequence $\{m_G^\nu(d\mathbf{k})\}_\nu$ of measures towards $m_G(d\mathbf{k})$ has been studied (see Lemma 3.2). The uncertain dimensionless spectrum, represented by $\tilde{\mu}_G^\nu(d\boldsymbol{\tau}; [y])$ for $[y]$ given in \mathcal{C}_y , is constructed from $\mu_G^\nu(d\boldsymbol{\tau}) = m_G^\nu(d\mathbf{k})$ and constitutes the uncertain spectral measure of random field $G^\nu(\cdot; \mathcal{S})$ given $\mathcal{S} = \{\boldsymbol{\omega}, [y]\} \in \mathcal{C}_\mathcal{S} = \mathcal{C}_\boldsymbol{\omega} \times \mathcal{C}_y$. Although a limit $G^\infty(\cdot; \mathcal{S})$ of random field $G^\nu(\cdot; \mathcal{S})$ exits for $\nu \rightarrow +\infty$ (see [22]), a convergence analysis is not useful for the probabilistic construction that is proposed because the limit is not given (unknown). The value of $\nu = (\nu_s)^3$ (see Definition 3.1) is chosen sufficiently large in order that Hypothesis 3 be verified. Proposition 3.2 shows that, for all $\mathcal{S} \in \mathcal{C}_\mathcal{S}$, the random field $G^\nu(\cdot; \mathcal{S})$, defined by Eq. (3.14), satisfies all the required properties (Gaussian, homogeneous, mean-square continuous, and normalization). We are therefore led to introduce the following definition in coherence with Proposition 2.1, Lemma 2.3, and Remark 2.4.

Definition 4.1 (*Random field $[\mathbf{C}(\cdot; \mathcal{S})]$ given \mathcal{S}*). We assume that ν is fixed and satisfies Hypothesis 3. The non-Gaussian random field $\mathbf{C}(\cdot; \mathcal{S})$ given $\mathcal{S} \in \mathcal{C}_\mathcal{S}$ is defined by Eq. (2.6) in which the 21 Gaussian random fields $\{G_{mn}(\mathbf{x}; \mathcal{S}), \mathbf{x} \in \mathbb{R}^3\}_{1 \leq m \leq n \leq 6}$ are replaced by 21 independent copies of the Gaussian real-valued random field $\{G^\nu(\mathbf{x}; \mathcal{S}), \mathbf{x} \in \mathbb{R}^3\}$ defined by Eq. (3.14), and denoted by $\{G_{mn}^\nu(\mathbf{x}; \mathcal{S}), \mathbf{x} \in \mathbb{R}^3\}_{1 \leq m \leq n \leq 6}$. For all $\mathcal{S} \in \mathcal{C}_\mathcal{S}$, $\mathbf{x} \in \mathbb{R}^3$, and for $1 \leq m \leq n \leq 6$, using Eq. (3.14) yields

$$(4.1) \quad G_{mn}^\nu(\mathbf{x}; \mathcal{S}) = g^\nu(\mathbf{x}; \mathcal{S}, \mathbf{Z}^{mn}, \boldsymbol{\Phi}^{mn}).$$

in which $\{\mathbf{Z}^{mn}, \boldsymbol{\Phi}^{mn}\}_{1 \leq m \leq n \leq 6}$ are 21 independent copies of \mathbb{R}^ν -valued random variables \mathbf{Z} and $\boldsymbol{\Phi}$ (see Definition 3.7), and we have

$$(4.2) \quad E\{G_{mn}^\nu(\mathbf{x}; \mathcal{S})\} = 0 \quad , \quad E\{G_{mn}^\nu(\mathbf{x}; \mathcal{S})^2\} = 1.$$

Proposition 4.1 (Properties of the non-Gaussian \mathbb{M}_6^+ -valued random field $\{[\mathbf{C}(\cdot; \mathcal{S})]\}$). *The non-Gaussian random field $[\mathbf{C}(\cdot; \mathcal{S})]$, defined in Definition 4.1 for a given uncertain spectral measure parameterized by \mathcal{S} , is a second-order random field such that*

$$(4.3) \quad \|[\mathbf{C}(\mathbf{x}; \mathcal{S})]\|_F \leq \Gamma_C \text{ a.s. } \quad , \quad \forall \mathbf{x} \in \mathbb{R}^3,$$

in which Γ_C is a second-order positive-valued random variable, independent of \mathbf{x} and \mathcal{S} , such that

$$(4.4) \quad E\{\Gamma_C^2\} = \gamma_{2,C}^2 < +\infty \quad , \quad E\{\Gamma_C^4\} = \gamma_{4,C}^4 < +\infty.$$

For all $\boldsymbol{\omega}$ and $\boldsymbol{\omega}'$ in \mathbb{R}^6 and for all \mathbf{x} in \mathbb{R}^3 ,

$$(4.5) \quad |([\mathbf{C}(\mathbf{x}; \mathcal{S})]\boldsymbol{\omega}, \boldsymbol{\omega}')_2| \leq \Gamma_C \|\boldsymbol{\omega}\|_2 \|\boldsymbol{\omega}'\|_2 \text{ a.s.}$$

Proof. (Proposition 4.1). For all $\mathbf{x} \in \mathbb{R}^3$ and $\mathbf{S} \in \mathcal{C}_S$, Proposition 2.1 and Definition 4.1 yield $[\mathbf{C}(\mathbf{x}; \mathbf{S})] = [\mathbf{L}(\mathbf{x}; \mathbf{S})]^T [\mathbf{L}(\mathbf{x}; \mathbf{S})]$ in which $[\mathbf{L}(\mathbf{x}; \mathbf{S})]_{mn} = \sigma_c G_{mn}^\nu(\mathbf{x}; \mathbf{S})$ for $1 \leq m < n \leq 6$ and $[\mathbf{L}(\mathbf{x}; \mathbf{S})]_{mm} = \sigma_c \sqrt{2h(G_{mm}^\nu(\mathbf{x}; \mathbf{S}); \alpha_m)}$ for $1 \leq m = n \leq 6$. We thus have, almost surely, $\|[\mathbf{C}(\mathbf{x}; \mathbf{S})]\|_F \leq \|[\mathbf{L}(\mathbf{x}; \mathbf{S})]\|_F^2 = \sum_m [\mathbf{L}(\mathbf{x}; \mathbf{S})]_{mm}^2 + \sum_{m < n} [\mathbf{L}(\mathbf{x}; \mathbf{S})]_{mn}^2 = \sigma_c^2 (\sum_m 2h(G_{mm}^\nu(\mathbf{x}; \mathbf{S}); \alpha_m) + \sum_{m < n} G_{mn}^\nu(\mathbf{x}; \mathbf{S})^2)$. Using Lemma 2.3-(i) yields

$$(4.6) \quad \|[\mathbf{C}(\mathbf{x}; \mathbf{S})]\|_F \leq \sigma_c^2 \left(4 \sum_m \alpha_m + 2 \sum_m G_{mm}^\nu(\mathbf{x}; \mathbf{S})^2 + \sum_{m < n} G_{mn}^\nu(\mathbf{x}; \mathbf{S})^2 \right).$$

For all $\mathbf{S} = \{\boldsymbol{\omega}, [y]\} \in \mathcal{C}_S = \mathcal{C}_\omega \times \mathcal{C}_y$, $\mathbf{z} \in \mathcal{C}_z$, and $\boldsymbol{\varphi} \in \mathcal{C}_\varphi$, Eq. (3.13) allows for writing $|g^\nu(\mathbf{x}; \mathbf{S}, \mathbf{z}, \boldsymbol{\varphi})| \leq \sum_{\beta \in \mathcal{B}} \sqrt{2\tilde{\chi}_\beta^\Delta([y])} z_\beta \leq \sqrt{2 \sum_{\beta \in \mathcal{B}} \tilde{\chi}_\beta^\Delta([y])} \sqrt{\sum_{\beta \in \mathcal{B}} z_\beta^2} \leq \sqrt{2 \sum_{\beta \in \mathcal{B}} z_\beta^2}$ because, from Proposition 3.1-(i), $\sum_{\beta \in \mathcal{B}} \tilde{\chi}_\beta^\Delta([y]) = 1$. For all $\mathbf{x} \in \mathbb{R}^3$, using Eq. (4.1) yields $G_{mn}^\nu(\mathbf{x}; \mathbf{S})^2 \leq 2 \sum_{\beta \in \mathcal{B}} (Z_\beta^{mn})^2$ almost surely. Using Eq. (4.6), we obtain Eq. (4.3) in which $\Gamma_C = \sigma_c^2 (4 \sum_m \alpha_m + 4 \sum_m \sum_{\beta \in \mathcal{B}} (Z_\beta^{mm})^2 + 2 \sum_{m < n} \sum_{\beta \in \mathcal{B}} (Z_\beta^{mn})^2)$ that is independent of \mathbf{x} and \mathbf{S} . The $21 \times \nu$ random variables $\{Z_\beta^{mn}\}_{mn, \beta}$ are independent copies of random variable $Z = \sqrt{-\log \Psi}$ whose probability measure is $P_Z(dz) = \mathbb{1}_{\mathbb{R}^+}(z) 2z \exp(-z^2) dz$. We have $E\{Z\} = \sqrt{\pi}/2$ and $E\{Z^2\} = 1$. Hence, $E\{\Gamma_C\} = \gamma_{1,C}$ with $\gamma_{1,C} = \sigma_c^2 (4 \sum_m \alpha_m + 54\nu) < +\infty$. For $p = 2$ or 4 , $E\{\Gamma_C^p\} = 2(2\sigma_c^2)^p \int_0^{+\infty} (2 \sum_m \alpha_m + 27\nu z^2)^p z \exp(-z^2) dz = \gamma_{p,C} < +\infty$ that yields Eq. (4.4) (note that $E\{\Gamma_C^4\} < +\infty$ implies $E\{\Gamma_C\} < +\infty$ and $E\{\Gamma_C^2\} < +\infty$). Since $E\{\|[\mathbf{C}(\mathbf{x}; \mathbf{S})]\|_F^2\} \leq E\{\Gamma_C^2\} < +\infty$, $[\mathbf{C}(\cdot; \mathbf{S})]$ is a second-order random field. Finally, for all $\mathbf{x} \in \mathbb{R}^3$, $\boldsymbol{\omega}$ and $\boldsymbol{\omega}'$ in \mathbb{R}^6 , $|\langle [\mathbf{C}(\mathbf{x}; \mathbf{S})] \boldsymbol{\omega}, \boldsymbol{\omega}' \rangle_2| \leq \|[\mathbf{C}(\mathbf{x}; \mathbf{S})]\|_F \|\boldsymbol{\omega}\|_2 \|\boldsymbol{\omega}'\|_2 \leq \|[\mathbf{C}(\mathbf{x}; \mathbf{S})]\|_F \|\boldsymbol{\omega}\|_2 \|\boldsymbol{\omega}'\|_2$, which yields Eq. (4.5) using Eq. (4.3). \square

Corollary 4.1 (Properties of the non-Gaussian \mathbb{M}_6^+ -valued random field $[\mathbb{C}(\cdot; \mathbf{S})]$). *Let $\mathbb{C}(\cdot; \mathbf{S})$ be the non-Gaussian second-order random field defined by Eq. (2.2) in which $\mathbb{C}(\cdot; \mathbf{S})$ satisfies the properties given in Proposition 4.1. For all $\boldsymbol{\omega} \in \mathbb{R}^6 \setminus \{0\}$,*

$$(4.7) \quad \frac{\langle [\mathbb{C}(\mathbf{x}; \mathbf{S})] \boldsymbol{\omega}, \boldsymbol{\omega} \rangle_2}{\|\boldsymbol{\omega}\|_2^2} \leq \|[\mathbb{C}(\mathbf{x}; \mathbf{S})]\|_F \leq \Gamma_{\mathbb{C}} \text{ a.s.}, \quad \forall \mathbf{x} \in \mathbb{R}^3,$$

with $\Gamma_{\mathbb{C}}$ a second-order \mathbb{R}^+ -valued random variable, independent of \mathbf{x} and \mathbf{S} , such that

$$(4.8) \quad E\{\Gamma_{\mathbb{C}}^2\} = \gamma_{2,\mathbb{C}}^2 < +\infty, \quad E\{\Gamma_{\mathbb{C}}^4\} = \gamma_{4,\mathbb{C}}^4 < +\infty.$$

For all $\boldsymbol{\omega}$ and $\boldsymbol{\omega}'$ in \mathbb{R}^6 and for all \mathbf{x} in \mathbb{R}^3 ,

$$(4.9) \quad |\langle [\mathbb{C}(\mathbf{x}; \mathbf{S})] \boldsymbol{\omega}, \boldsymbol{\omega}' \rangle_2| \leq \Gamma_{\mathbb{C}} \|\boldsymbol{\omega}\|_2 \|\boldsymbol{\omega}'\|_2 \text{ a.s.},$$

$$(4.10) \quad \langle [\mathbb{C}(\mathbf{x}; \mathbf{S})] \boldsymbol{\omega}, \boldsymbol{\omega} \rangle_2 \geq \underline{c}_\epsilon \|\boldsymbol{\omega}\|_2^2 \text{ a.s.}$$

in which $\underline{c}_\epsilon = \underline{c}_0 \epsilon / (1 + \epsilon)$ is a finite positive constant independent of \mathbf{x} and \mathbf{S} .

Proof. (Corollary 4.1). Equation (2.2) yields $\|[\mathbb{C}(\mathbf{x}; \mathbf{S})]\|_F \leq (1 + \epsilon)^{-1} \|[\mathbb{L}]^T\|_F \|[\mathbb{L}]\|_F$ ($\epsilon \|[\mathbb{I}_6]\|_F + \|[\mathbf{C}(\mathbf{x}; \mathbf{S})]\|_F$). We have $\|[\mathbb{L}]^T\|_F = \|[\mathbb{L}]\|_F = (\text{tr}[\mathbb{C}])^{1/2}$ and $\|[\mathbb{I}_6]\|_F = \sqrt{6}$. Eqs. (2.1) and (4.3) yield $\|[\mathbb{C}(\mathbf{x}; \mathbf{S})]\|_F \leq \underline{c}_1 (1 + \epsilon)^{-1} (\epsilon \sqrt{6} + \|[\mathbf{C}(\mathbf{x}; \mathbf{S})]\|_F) \leq \Gamma_{\mathbb{C}}$ almost surely with $\Gamma_{\mathbb{C}} = \underline{c}_1 (1 + \epsilon)^{-1} (\epsilon \sqrt{6} + \Gamma_C)$, which is the second part of Eq. (4.7). For all $\boldsymbol{\omega}$ and $\boldsymbol{\omega}'$ in \mathbb{R}^6 , we have $|\langle [\mathbb{C}(\mathbf{x}; \mathbf{S})] \boldsymbol{\omega}, \boldsymbol{\omega}' \rangle_2| \leq \|[\mathbb{C}(\mathbf{x}; \mathbf{S})]\|_F \|\boldsymbol{\omega}\|_2 \|\boldsymbol{\omega}'\|_2$. Taking $\boldsymbol{\omega}' = \boldsymbol{\omega} \in \mathbb{R}^6 \setminus \{0\}$ yields the first part of Eq. (4.7). Then using the second part of Eq. (4.7) yields Eq. (4.9). From Eq. (2.2), it can be deduced that $\langle [\mathbb{C}(\mathbf{x}; \mathbf{S})] \boldsymbol{\omega}, \boldsymbol{\omega} \rangle_2 = (\epsilon \langle [\mathbb{C}] \boldsymbol{\omega}, \boldsymbol{\omega} \rangle_2 + \langle [\mathbf{C}(\mathbf{x}; \mathbf{S})] [\mathbb{L}] \boldsymbol{\omega}, [\mathbb{L}] \boldsymbol{\omega} \rangle_2) / (1 + \epsilon)$. From Eq. (2.1) and since $[\mathbf{C}(\mathbf{x}; \mathbf{S})]$ is a \mathbb{M}_6^+ -valued random variable, we obtain Eq. (4.10). \square

Definition 4.2 (*Random field $[\tilde{\mathbb{C}}]$ with uncertain spectral measure*). We assume that ν is fixed and satisfies Hypothesis 3. The random field $\{[\tilde{\mathbb{C}}(\mathbf{x})] \in \mathbb{R}^3\}$ with uncertain spectral measure is defined by

$$(4.11) \quad [\tilde{\mathbb{C}}(\mathbf{x})] = [\mathbb{C}(\mathbf{x}; \mathbf{S})] \quad , \quad \forall \mathbf{x} \in \mathbb{R}^3 ,$$

in which $\mathbf{S} = \{\mathbf{W}, [\mathbf{Y}]\}$ is the $\mathbb{R}^3 \times \mathbb{M}_{3, \hat{\nu}_s}$ -valued random variable defined by Definition 3.6.

Remark 4.3. Random field $\{[\tilde{\mathbb{C}}(\mathbf{x})] \in \mathbb{R}^3\}$ defined by Eq. (4.11) is the random field with uncertain spectral measure. As explained in Remark 2.4, $E\{[\mathbb{C}(\mathbf{x}; \mathbf{S})]\} = [\mathbb{C}]$ for all $\mathbf{x} \in \mathbb{R}^3$, we have not $E\{[\tilde{\mathbb{C}}(\mathbf{x})]\} = [\mathbb{C}]$, but we have the approximation $E\{[\tilde{\mathbb{C}}(\mathbf{x})]\} \simeq [\mathbb{C}]$.

5. STOCHASTIC ELLIPTIC BOUNDARY VALUE PROBLEM FOR STOCHASTIC HOMOGENIZATION

We consider a heterogeneous complex elastic microstructure occupying domain Ω , which by definition is a microstructure that cannot be described in terms of its constituents at the microscale. This is typically the case of live tissues. In such a case, the stochastic model of the apparent elasticity field can be constructed at the mesoscale that corresponds to the scale of the spatial correlation length of the microstructure Ω as proposed in [29, 30, 32]. The stochastic homogenization from the mesoscale to the macroscale allows the effective elasticity tensor $\tilde{\mathbb{C}}^{\text{eff}}$ to be constructed. The study of the statistical properties of $\tilde{\mathbb{C}}^{\text{eff}}$ allows for analyzing the scale separation. The separation is obtained if the statistical fluctuations of $\tilde{\mathbb{C}}^{\text{eff}}$ are sufficiently small and, in this case, Ω is a representative volume element (RVE) [11, 20, 21, 30]. Such a separation occurs if the spatial correlation length at the mesoscale is sufficiently small with respect to the characteristic geometrical dimension of Ω . If not, $\tilde{\mathbb{C}}^{\text{eff}}$ exhibits significant statistical fluctuations and therefore, Ω is not a RVE.

The deterministic part of the formulation used in Sections 5.1 to 5.3 to write the problem of homogenization on Ω is that proposed in [2] for homogeneous deformations on the boundary $\partial\Omega$. We use the convention for summations over repeated Latin indices j , p , and q taking values in $\{1, 2, 3\}$.

5.1. Definition of the stochastic boundary value problem (BVP). Let Ω be a bounded open subset of \mathbb{R}^3 with a sufficiently regular boundary $\partial\Omega$. Let \mathbf{S} be fixed in $\mathcal{C}_{\mathcal{S}}$. For all ℓ and r in $\{1, 2, 3\}$, we have to find the \mathbb{R}^3 -valued random field $\{\tilde{\mathbf{U}}^{\ell r}(\mathbf{x}) = (\tilde{U}_1^{\ell r}(\mathbf{x}), \tilde{U}_2^{\ell r}(\mathbf{x}), \tilde{U}_3^{\ell r}(\mathbf{x})), \mathbf{x} \in \bar{\Omega}\}$, defined on $(\Theta, \mathcal{T}, \mathcal{P})$, indexed by $\bar{\Omega}$, such that almost surely,

$$(5.1) \quad -\frac{\partial}{\partial x_j} (\mathbb{C}_{ijpq}(\mathbf{x}; \mathbf{S}) \varepsilon_{pq}(\tilde{\mathbf{U}}^{\ell r}(\mathbf{x}))) = 0 \quad , \quad \forall \mathbf{x} \in \Omega \quad , \quad i = 1, 2, 3 ,$$

$$(5.2) \quad \tilde{\mathbf{U}}^{\ell r}(\mathbf{x}) = \tilde{\mathbf{u}}_0^{\ell r}(\mathbf{x}) \quad , \quad \forall \mathbf{x} \in \partial\Omega ,$$

in which $\varepsilon_{pq}(\mathbf{u}) = (\partial u_p / \partial x_q + \partial u_q / \partial x_p) / 2$ for $\mathbf{u} = (u_1, u_2, u_3)$ and where for all $\mathbf{x} \in \partial\Omega$, $\tilde{\mathbf{u}}_0^{\ell r}(\mathbf{x}) = (\tilde{u}_{0,1}^{\ell r}(\mathbf{x}), \tilde{u}_{0,2}^{\ell r}(\mathbf{x}), \tilde{u}_{0,3}^{\ell r}(\mathbf{x}))$ is defined by $\tilde{u}_{0,j}^{\ell r}(\mathbf{x}) = (\delta_{j\ell} x_r + \delta_{jr} x_\ell) / 2$ with $\delta_{j\ell}$ the Kronecker symbol. The fourth-order tensor-valued random field $\{\mathbb{C}_{ijpq}(\cdot; \mathbf{S})\}_{ijpq}$ is such that $\mathbb{C}_{ijpq} = \mathbb{C}_{jipq} = \mathbb{C}_{ijqp} = \mathbb{C}_{pqij}$ for i, j, p , and q in $\{1, 2, 3\}$ and is such that $\mathbb{C}_{ijpq}(\cdot; \mathbf{S}) = [\mathbb{C}(\cdot, \mathbf{S})]_{\mathbf{ij}}$ in which $\mathbf{i} = (i, j)$ with $1 \leq i \leq j \leq 3$ and $\mathbf{j} = (p, q)$ with $1 \leq p \leq q \leq 3$ are indices with values in $\{1, \dots, 6\}$, and where the \mathbb{M}_6^+ -valued random field $[\mathbb{C}(\cdot; \mathbf{S})]$ is the one constructed in Section 4 and whose properties are given by Corollary 4.1.

5.2. Random effective tensor from stochastic homogenization and its random eigenvalues. For \mathcal{S} fixed in \mathcal{C}_S , for i, j, ℓ , and r in $\{1, 2, 3\}$ the component $\mathbb{C}_{ij\ell r}^{\text{eff}}(\mathcal{S})$ of the random fourth-order effective tensor $\mathbb{C}^{\text{eff}}(\mathcal{S})$ is defined by

$$(5.3) \quad \mathbb{C}_{ij\ell r}^{\text{eff}}(\mathcal{S}) = \frac{1}{|\Omega|} \int_{\Omega} \mathbb{C}_{ijpq}(\mathbf{x}; \mathcal{S}) \varepsilon_{pq}(\tilde{\mathbf{U}}^{\ell r}(\mathbf{x})) d\mathbf{x},$$

in which $\tilde{\mathbf{U}}^{\ell r}$ is the \mathbb{R}^3 -valued random field that satisfies Eqs. (5.1) and (5.2) and where $|\Omega| = \int_{\Omega} d\mathbf{x}$. The fourth-order effective tensor $\mathbb{C}^{\text{eff}}(\mathcal{S})$ satisfies the symmetry and positive-definiteness properties [2]. We can thus define the effective \mathbb{M}_6^+ -valued random matrix $[\mathbb{C}^{\text{eff}}(\mathcal{S})]$ associated with random tensor $\mathbb{C}^{\text{eff}}(\mathcal{S})$, which is such that $[\mathbb{C}^{\text{eff}}(\mathcal{S})]_{\mathbf{ij}} = \mathbb{C}_{ij\ell r}^{\text{eff}}(\mathcal{S})$ in which $\mathbf{i} = (i, j)$ with $1 \leq i \leq j \leq 3$ and $\mathbf{j} = (\ell, r)$ with $1 \leq \ell \leq r \leq 3$.

5.3. Transforming the nonhomogeneous Dirichlet BVP in a homogeneous Dirichlet BVP. For fixed ℓ and r , since $\mathbf{x} \mapsto \tilde{\mathbf{u}}_0^{\ell r}(\mathbf{x})$ is a linear function in \mathbf{x} , we can perform the following translation (without having to resort the trace theorem in Hilbert spaces),

$$(5.4) \quad \tilde{\mathbf{U}}^{\ell r}(\mathbf{x}) = \mathbf{U}^{\ell r}(\mathbf{x}) + \tilde{\mathbf{u}}_0^{\ell r}(\mathbf{x}) \quad , \quad \forall \mathbf{x} \in \bar{\Omega}.$$

Since, $\forall \mathbf{x} \in \bar{\Omega}$, $\varepsilon_{pq}(\tilde{\mathbf{u}}_0^{\ell r}(\mathbf{x})) = (\delta_{p\ell}\delta_{qr} + \delta_{pr}\delta_{q\ell})/2$ and $\mathbb{C}_{ij\ell r}(\mathbf{x}; \mathcal{S}) = \mathbb{C}_{ijr\ell}(\mathbf{x}; \mathcal{S})$, for all ℓ and r in $\{1, 2, 3\}$, the nonhomogeneous Dirichlet BVP defined by Eqs. (5.1) and (5.2) becomes the following homogeneous Dirichlet BVP for the \mathbb{R}^3 -valued random field $\{\mathbf{U}^{\ell r}(\mathbf{x}) = (U_1^{\ell r}(\mathbf{x}), U_2^{\ell r}(\mathbf{x}), U_3^{\ell r}(\mathbf{x})), \mathbf{x} \in \bar{\Omega}\}$, defined on $(\Theta, \mathcal{T}, \mathcal{P})$, indexed by $\bar{\Omega}$, such that almost surely,

$$(5.5) \quad -\frac{\partial}{\partial x_j}(\mathbb{C}_{ijpq}(\mathbf{x}; \mathcal{S}) \varepsilon_{pq}(\mathbf{U}^{\ell r}(\mathbf{x}))) = f_i^{\ell r}(\mathbf{x}; \mathcal{S}) \quad , \quad \forall \mathbf{x} \in \Omega \quad , \quad i = 1, 2, 3,$$

$$(5.6) \quad \mathbf{U}^{\ell r}(\mathbf{x}) = \mathbf{0} \quad , \quad \forall \mathbf{x} \in \partial\Omega,$$

in which $f_i^{\ell r}(\mathbf{x}; \mathcal{S}) = \frac{\partial}{\partial x_j}(\mathbb{C}_{ij\ell r}(\mathbf{x}; \mathcal{S}))$ that, from Proposition 2.1 and Eqs. (3.13), (3.14), and (4.1), exists almost surely.

5.4. Analysis of the stochastic homogeneous Dirichlet BVP. (i) *Definition of random vector Ξ .* Let $\Xi = \{\{\mathbf{Z}^{mn}, 1 \leq m \leq n \leq 6\}, \{\Phi^{mn}, 1 \leq m \leq n \leq 6\}\}$ be the second-order random variable on $(\Theta, \mathcal{T}, \mathcal{P})$ with values in \mathbb{R}^{n_ξ} with $n_\xi = 2 \times 21 \times \nu$, whose probability measure is $P_\Xi = (\otimes_{m,n} P_{\mathbf{Z}^{mn}}) \otimes (\otimes_{m,n} P_{\Phi^{mn}})$ in which $P_{\mathbf{Z}^{mn}} = P_{\mathbf{Z}}$ and $P_{\Phi^{mn}} = P_{\Phi}$ for all $1 \leq m \leq n \leq 6$ (see Definitions 3.7 and 4.1). Let $\mathcal{C}_\xi \subset \mathbb{R}^{n_\xi}$ be the support of P_Ξ , which is known and can easily be written. Consequently, we have $E\{\|\Xi\|_2^2\} = \int_{\mathbb{R}^{n_\xi}} \|\xi\|_2^2 P_\Xi(d\xi) = \int_{\mathcal{C}_\xi} \|\xi\|_2^2 P_\Xi(d\xi) < +\infty$.

(ii) *Definition of mappings $\xi \mapsto \mathbb{C}(\cdot; \mathcal{S}, \xi)$ and $\xi \mapsto \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi)$.* For \mathcal{S} fixed in \mathcal{C}_S , the fourth-order tensor-valued random field $\mathbb{C}(\cdot; \mathcal{S})$ is written as $\mathbb{C}_{ijpq}(\mathbf{x}; \mathcal{S}) = [\mathbb{C}(\mathbf{x}; \mathcal{S})]_{\mathbf{ij}}$ in which $\mathbf{i} = (i, j)$ with $i \leq j$ and $\mathbf{j} = (p, q)$ with $p \leq q$. Taking into account the construction presented in Sections 2 to 4, the random field $[\mathbb{C}(\cdot; \mathcal{S})]$ is defined by a \mathbb{M}_6^+ -valued measurable mapping $\xi \mapsto [\mathbb{C}(\cdot; \mathcal{S}, \xi)]$ on \mathcal{C}_ξ such that $[\mathbb{C}(\cdot; \mathcal{S})] = [\mathbb{C}(\cdot; \mathcal{S}, \Xi)]$. Therefore, the fourth-order random field $\mathbb{C}(\cdot; \mathcal{S})$ is defined by a measurable mapping $\xi \mapsto \mathbb{C}(\cdot; \mathcal{S}, \xi)$ on \mathcal{C}_ξ such that

$$(5.7) \quad \mathbb{C}_{ijpq}(\cdot; \mathcal{S}) = \mathbb{C}_{ijpq}(\cdot; \mathcal{S}, \Xi) \quad , \quad i, j, p, q \in \{1, 2, 3\}.$$

Similarly, random field $\mathbf{U}^{\ell r}$ involved in the stochastic BVP, defined by Eqs. (5.5) and (5.6), depends only on \mathcal{S} and Ξ , and is defined by a measurable mapping $\xi \mapsto \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi)$ on \mathcal{C}_ξ such that $\mathbf{U}^{\ell r} = \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \Xi)$ for ℓ and r in $\{1, 2, 3\}$.

(iii) *Definition of Hilbert space \mathbb{H} .* Let $\mathbb{H} = \{\mathbf{v} = (v_1, v_2, v_3); v_j \in H^1(\Omega) \text{ for } j = 1, 2, 3; \mathbf{v} = \mathbf{0} \text{ on } \partial\Omega\}$ be the Hilbert space equipped with the inner product and the associated norm,

$$(5.8) \quad \langle \mathbf{u}, \mathbf{v} \rangle_{\mathbb{H}} = \int_{\Omega} \varepsilon_{pq}(\mathbf{u}(\mathbf{x})) \varepsilon_{pq}(\mathbf{v}(\mathbf{x})) d\mathbf{x} \quad , \quad \|\mathbf{v}\|_{\mathbb{H}} = (\langle \mathbf{v}, \mathbf{v} \rangle_{\mathbb{H}})^{1/2} .$$

Note that $\|\mathbf{v}\|_{\mathbb{H}}$ is a norm on \mathbb{H} due to the Korn inequality and because $\mathbf{v} = \mathbf{0}$ on $\partial\Omega$ (see for instance [5]). Introducing the matrix $[\varepsilon]$ such that $[\varepsilon]_{pq} = \varepsilon_{pq}$, Eq. (5.8) can be rewritten as

$$(5.9) \quad \langle \mathbf{u}, \mathbf{v} \rangle_{\mathbb{H}} = \int_{\Omega} \langle [\varepsilon(\mathbf{u}(\mathbf{x}))], [\varepsilon(\mathbf{v}(\mathbf{x}))] \rangle_F d\mathbf{x} \quad , \quad \|\mathbf{v}\|_{\mathbb{H}}^2 = \int_{\Omega} \|[\varepsilon(\mathbf{v}(\mathbf{x}))]\|_F^2 d\mathbf{x} .$$

(iv) *Definition of bilinear form $b(\cdot, \cdot; \mathcal{S}, \boldsymbol{\xi})$ and linear form $\mathcal{L}^{\ell r}(\cdot; \mathcal{S}, \boldsymbol{\xi})$.* For \mathcal{S} fixed in $\mathcal{C}_{\mathcal{S}}$, for all $\boldsymbol{\xi} \in \mathcal{C}_{\boldsymbol{\xi}}$, for ℓ and r in $\{1, 2, 3\}$, and using Eq. (5.7), we define the bilinear form $(\mathbf{u}, \mathbf{v}) \mapsto b(\mathbf{u}, \mathbf{v}; \mathcal{S}, \boldsymbol{\xi}) : \mathbb{H} \times \mathbb{H} \rightarrow \mathbb{R}$ such that

$$(5.10) \quad b(\mathbf{u}, \mathbf{v}; \mathcal{S}, \boldsymbol{\xi}) = \int_{\Omega} \mathfrak{c}_{ijpq}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi}) \varepsilon_{pq}(\mathbf{u}(\mathbf{x})) \varepsilon_{ij}(\mathbf{v}(\mathbf{x})) d\mathbf{x} ,$$

and the linear form $\mathbf{v} \mapsto \mathcal{L}^{\ell r}(\mathbf{v}; \mathcal{S}, \boldsymbol{\xi}) : \mathbb{H} \rightarrow \mathbb{R}$ such that

$$(5.11) \quad \mathcal{L}^{\ell r}(\mathbf{v}; \mathcal{S}, \boldsymbol{\xi}) = - \int_{\Omega} \mathfrak{c}_{\ell r ij}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi}) \varepsilon_{ij}(\mathbf{v}(\mathbf{x})) d\mathbf{x} ,$$

whose right-hand side member in Eq. (5.11) is the transformation of $\int_{\Omega} f_i^{\ell r}(\mathbf{x}; \mathcal{S}) v_i(\mathbf{x}) d\mathbf{x}$ for which we have used $\mathbf{v} = \mathbf{0}$ on $\partial\Omega$ and the symmetry property $\mathfrak{c}_{ij\ell r} = \mathfrak{c}_{j\ell ir} = \mathfrak{c}_{\ell r ij}$.

(v) *Type of stochastic solution sought.* From a computational point of view, a solution of the stochastic BVP defined by Eqs. (5.5) and (5.6) will be constructed by using the Monte Carlo simulation method. Consequently, we only need to analyze the strong stochastic solution of the weak formulation of this stochastic BVP and the weak stochastic solution is not useful. We then limit Proposition 5.1 to the strong stochastic solution.

Proposition 5.1 (Weak formulation of the stochastic homogeneous Dirichlet BVP and its strong stochastic solution). *(i) For \mathcal{S} fixed in $\mathcal{C}_{\mathcal{S}}$ and for $1 \leq \ell \leq r \leq 3$, the weak formulation of the stochastic BVP defined by Eqs. (5.5) and (5.6) is: for P_{Ξ} -almost all $\boldsymbol{\xi}$ in $\mathcal{C}_{\boldsymbol{\xi}} \subset \mathbb{R}^{n_{\boldsymbol{\xi}}}$, find $\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \boldsymbol{\xi})$ in \mathbb{H} such that*

$$(5.12) \quad b(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \boldsymbol{\xi}), \mathbf{v}; \mathcal{S}, \boldsymbol{\xi}) = \mathcal{L}^{\ell r}(\mathbf{v}; \mathcal{S}, \boldsymbol{\xi}) \quad , \quad \forall \mathbf{v} \in \mathbb{H} .$$

(ii) For $1 \leq \ell \leq r \leq 3$, there exists a unique solution $\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \boldsymbol{\xi}) \in \mathbb{H}$ (strong stochastic solution) such that Eq. (5.12) holds and $\mathbf{u}^{r\ell}(\cdot; \mathcal{S}, \boldsymbol{\xi}) = \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \boldsymbol{\xi})$.

(iii) The associated stochastic solution $\mathbf{U}^{\ell r}(\cdot; \mathcal{S}) = \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \boldsymbol{\Xi})$ is of second-order,

$$(5.13) \quad E\{\|\mathbf{U}^{\ell r}(\cdot; \mathcal{S})\|_{\mathbb{H}}^2\} = \gamma_u^2 < +\infty .$$

Proof. (Proposition 5.1).

(i) Using Eqs. (5.7) and (5.8) to (5.11), it is easy to prove that Eq. (5.12) is the weak formulation of Eqs. (5.5) and (5.6).

(ii) Let $[\mathfrak{c}^{\ell r}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})]$ be the (3×3) real matrix such that $[\mathfrak{c}^{\ell r}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})]_{ij} = \mathfrak{c}_{\ell r ij}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})$. Therefore Eq. (5.11) can be rewritten as $\mathcal{L}^{\ell r}(\mathbf{v}; \mathcal{S}, \boldsymbol{\xi}) = - \int_{\Omega} \langle [\mathfrak{c}^{\ell r}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})], [\varepsilon(\mathbf{v}(\mathbf{x}))] \rangle_F d\mathbf{x}$ and using Eq. (5.9) yield $|\mathcal{L}^{\ell r}(\mathbf{v}; \mathcal{S}, \boldsymbol{\xi})| \leq (\int_{\Omega} \|[\mathfrak{c}^{\ell r}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})]\|_F^2 d\mathbf{x})^{1/2} \|\mathbf{v}\|_{\mathbb{H}}$. For all ℓ and r in $\{1, 2, 3\}$ and since $\mathfrak{c}_{\ell r ij} = \mathfrak{c}_{\ell r ji}$, we have $\|[\mathfrak{c}^{\ell r}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})]\|_F^2 = \sum_{i,j} \mathfrak{c}_{\ell r ij}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})^2 \leq 2 \sum_{\ell' \leq r', i \leq j} \mathfrak{c}_{\ell' r' ij}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})^2 = 2 \|[\mathfrak{c}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})]\|_F^2$ in which we have used the notation $[\mathfrak{c}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})]_{\mathbf{j}\mathbf{i}} = \mathfrak{c}_{\ell' r' ij}(\mathbf{x}; \mathcal{S}, \boldsymbol{\xi})$ in which $\mathbf{j}' = (\ell', r')$ with $\ell' \leq r'$ and $\mathbf{i} = (i, j)$ with

$i \leq j$. Taking into account Proposition 4.1 and its proof, Eqs. (4.7) and (4.8) of Corollary 4.1 with its proof, show that $\Gamma_{\mathbb{C}} = \gamma_{\mathbb{C}}(\Xi)$ in which $\xi \mapsto \gamma_{\mathbb{C}}(\xi)$ is a positive-valued measurable mapping on \mathcal{C}_{ξ} , which is independent of \mathcal{S} and such that

$$(5.14) \quad E\{\Gamma_{\mathbb{C}}^2\} = \int_{\mathcal{C}_{\xi}} \gamma_{\mathbb{C}}(\xi)^2 P_{\Xi}(d\xi) = \underline{\gamma}_{\mathbb{C}}^2 < +\infty.$$

Equation (4.7) shows that

$$(5.15) \quad \|\mathbb{C}(\mathbf{x}; \mathcal{S}, \xi)\|_F \leq \gamma_{\mathbb{C}}(\xi), \text{ for } P_{\Xi} \text{- almost all } \xi \text{ in } \mathcal{C}_{\xi}.$$

It can then be deduced that

$$(5.16) \quad |\mathcal{L}^{\ell r}(\mathbf{v}; \mathcal{S}, \xi)| \leq \sqrt{2} |\Omega|^{1/2} \gamma_{\mathbb{C}}(\xi) \|\mathbf{v}\|_{\mathbb{H}},$$

which shows that linear form $\mathbf{v} \mapsto \mathcal{L}^{\ell r}(\mathbf{v}; \mathcal{S}, \xi)$ is continuous on \mathbb{H} for P_{Ξ} -almost ξ in \mathcal{C}_{ξ} . Equation (4.9) shows that, $\forall \boldsymbol{\omega}$ and $\boldsymbol{\omega}'$ in \mathbb{R}^6 , $|\langle \mathbb{C}(\mathbf{x}; \mathcal{S}, \xi) \boldsymbol{\omega}, \boldsymbol{\omega}' \rangle_2| \leq \gamma_{\mathbb{C}}(\xi) \|\boldsymbol{\omega}\|_2 \|\boldsymbol{\omega}'\|_2$ for P_{Ξ} -almost ξ in \mathcal{C}_{ξ} . Using Eq. (5.10) and taking into account the symmetry properties of ε_{pq} and \mathbb{C}_{ijpq} yield, for all \mathbf{u} and \mathbf{v} in \mathbb{H} , $|b(\mathbf{u}, \mathbf{v}; \mathcal{S}, \xi)| \leq 2 \gamma_{\mathbb{C}}(\xi) \|\mathbf{u}\|_{\mathbb{H}} \|\mathbf{v}\|_{\mathbb{H}}$, which shows that bilinear form $(\mathbf{u}, \mathbf{v}) \mapsto b(\mathbf{u}, \mathbf{v}; \mathcal{S}, \xi)$ is continuous on $\mathbb{H} \times \mathbb{H}$ for P_{Ξ} -almost ξ in \mathcal{C}_{ξ} . Equation (4.10) shows that, $\forall \boldsymbol{\omega} \in \mathbb{R}^6$, $\langle \mathbb{C}(\mathbf{x}; \mathcal{S}, \xi) \boldsymbol{\omega}, \boldsymbol{\omega} \rangle_2 \geq \underline{c}_{\varepsilon} \|\boldsymbol{\omega}\|_2^2$ for P_{Ξ} -almost ξ in \mathcal{C}_{ξ} . From Eq. (5.10), it can be deduced that, $\forall \mathbf{v} \in \mathbb{H}$,

$$(5.17) \quad b(\mathbf{v}, \mathbf{v}; \mathcal{S}, \xi) \geq \underline{c}_{\varepsilon} \|\mathbf{v}\|_{\mathbb{H}}^2,$$

which proves that bilinear form $(\mathbf{u}, \mathbf{v}) \mapsto b(\mathbf{u}, \mathbf{v}; \mathcal{S}, \xi)$ is coercive for P_{Ξ} -almost ξ in \mathcal{C}_{ξ} . Due to the continuity and coercivity of bilinear form $b(\cdot, \cdot; \mathcal{S}, \xi)$ and due to the continuity of linear form $\mathcal{L}^{\ell r}(\cdot; \mathcal{S}, \xi)$ for P_{Ξ} -almost ξ in \mathcal{C}_{ξ} , the use of the Lax-Milgram theorem [13, 15] allows for proving (ii) of the Proposition.

(iii) Taking $\mathbf{v} = \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi)$ in Eq. (5.12) yields $b(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi), \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi); \mathcal{S}, \xi) = |\mathcal{L}^{\ell r}(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi); \mathcal{S}, \xi)|$. From Eq. (5.16), it can be deduced that $|\mathcal{L}^{\ell r}(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi); \mathcal{S}, \xi)| \leq \sqrt{2} |\Omega|^{1/2} \gamma_{\mathbb{C}}(\xi) \|\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi)\|_{\mathbb{H}}$ and using Eq. (5.17) yield $\underline{c}_{\varepsilon} \|\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi)\|_{\mathbb{H}}^2 \leq b(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi), \mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi); \mathcal{S}, \xi)$. Consequently, we obtain

$$(5.18) \quad \|\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi)\|_{\mathbb{H}} \leq \frac{\sqrt{2} |\Omega|^{1/2}}{\underline{c}_{\varepsilon}} \gamma_{\mathbb{C}}(\xi).$$

Finally, $E\{\|\mathbf{U}^{\ell r}(\cdot; \mathcal{S})\|_{\mathbb{H}}^2\} = \int_{\mathcal{C}_{\xi}} \|\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi)\|_{\mathbb{H}}^2 P_{\Xi}(d\xi) \leq 2 |\Omega| \underline{c}_{\varepsilon}^{-2} \int_{\mathcal{C}_{\xi}} \gamma_{\mathbb{C}}(\xi)^2 P_{\Xi}(d\xi)$ and using Eq. (5.14) yield $E\{\|\mathbf{U}^{\ell r}(\cdot; \mathcal{S})\|_{\mathbb{H}}^2\} = 2 |\Omega| \underline{\gamma}_{\mathbb{C}}^2 / \underline{c}_{\varepsilon}^2$ that is Eq. (5.13) with $\underline{\gamma}_u^2 = 2 |\Omega| \underline{\gamma}_{\mathbb{C}}^2 / \underline{c}_{\varepsilon}^2$. \square

5.5. Random eigenvalues of the random effective elasticity matrix. For $\mathcal{S} = \{\mathbf{w}, [y]\} \in \mathcal{C}_{\mathcal{S}} = \mathcal{C}_w \times \mathcal{C}_y$, the random effective elasticity matrix $[\mathbb{C}^{\text{eff}}(\mathcal{S})]$ defined in Section 5.2 can be written as $[\mathbb{C}^{\text{eff}}(\mathcal{S})] = [\mathbb{C}^{\text{eff}}(\mathcal{S}, \Xi)]$ in which $\xi \mapsto [\mathbb{C}^{\text{eff}}(\mathcal{S}, \xi)]$ is a \mathbb{M}_6^+ -valued measurable mapping on $\mathcal{C}_{\xi} \subset \mathbb{R}^{n_{\xi}}$. For all $\mathcal{S} \in \mathcal{C}_{\mathcal{S}}$ and $\xi \in \mathcal{C}_{\xi}$, let $\lambda_1(\mathcal{S}, \xi) \geq \dots \geq \lambda_6(\mathcal{S}, \xi) > 0$ be the eigenvalues of matrix $[\mathbb{C}^{\text{eff}}(\mathcal{S}, \xi)] \in \mathbb{M}_6^+$ and let $\boldsymbol{\lambda}(\mathcal{S}, \xi) = (\lambda_1(\mathcal{S}, \xi), \dots, \lambda_6(\mathcal{S}, \xi)) \in (\mathbb{R}^{+*})^6$. Let $\mathcal{S} = \{\mathbf{W}, [\mathbf{Y}]\}$ be the $\mathbb{R}^3 \times \mathbb{M}_{3, \hat{\nu}_s}$ -valued random variable defined in Definition 3.6 and Proposition 3.1-(ii), for which the support of its probability measure $P_{\mathcal{S}} = P_{\mathbf{W}}(d\mathbf{w}) \otimes P_{[\mathbf{Y}]}(d\mathbf{y})$ is subset $\mathcal{C}_{\mathcal{S}} = \mathcal{C}_w \times \mathcal{C}_y$ of $\mathbb{R}^3 \times \mathbb{M}_{3, \hat{\nu}_s}$. The random effective elasticity matrix $[\tilde{\mathbb{C}}^{\text{eff}}]$, which corresponds to the elasticity random field $\tilde{\mathbb{C}}$ for which its spectral measure is uncertain, can then be written as $[\tilde{\mathbb{C}}^{\text{eff}}] = [\mathbb{C}^{\text{eff}}(\mathcal{S})] = [\mathbb{C}^{\text{eff}}(\mathcal{S}, \Xi)]$. Let $\{\tilde{\Lambda}_j = \lambda_j(\mathcal{S}, \Xi), j = 1, \dots, 6\}$ be the ordered (a.s) random eigenvalues of $[\tilde{\mathbb{C}}^{\text{eff}}]$. Let $\tilde{\boldsymbol{\Lambda}} = (\tilde{\Lambda}_1, \dots, \tilde{\Lambda}_6)$ be the \mathbb{R}^6 -valued random variable whose support of its probability measure $P_{\tilde{\boldsymbol{\Lambda}}}(d\tilde{\boldsymbol{\Lambda}})$ is $(\mathbb{R}^{+*})^6$. The operator norm of $[\tilde{\mathbb{C}}^{\text{eff}}]$ is $\|[\tilde{\mathbb{C}}^{\text{eff}}]\|_2 = \tilde{\Lambda}_1$.

Corollary 5.1 (Second-order properties of the random eigenvalues). *Under proposition 5.1, $\tilde{\Lambda}$ is a second-order \mathbb{R}^6 -valued random variable,*

$$(5.19) \quad E\{\|\tilde{\Lambda}\|_2^2\} = \gamma_\lambda^2 < +\infty.$$

Proof. (Corollary 5.1). From Eqs. (5.3), (5.4), and (5.11), it can be deduced that $\mathfrak{c}_{ij\ell r}^{\text{eff}}(\mathcal{S}, \xi) = -|\Omega|^{-1} \mathcal{L}^{ij}(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi); \mathcal{S}, \xi) + |\Omega|^{-1} \int_{\Omega} \mathfrak{c}_{ij\ell r}(\mathbf{x}; \mathcal{S}, \xi) d\mathbf{x}$. Since $(a+b)^2 \leq 2(a^2+b^2)$, we have $\|\boldsymbol{\lambda}(\mathcal{S}, \xi)\|_2^2 = \|\mathfrak{c}^{\text{eff}}(\mathcal{S}, \xi)\|_F^2 = \sum_{i \leq j, \ell \leq r} \mathfrak{c}_{ij\ell r}^{\text{eff}}(\mathcal{S}, \xi)^2 \leq 2|\Omega|^{-2} \sum_{i \leq j, \ell \leq r} \{\mathcal{L}^{ij}(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi); \mathcal{S}, \xi)\|^2 + (\int_{\Omega} \mathfrak{c}_{ij\ell r}(\mathbf{x}; \mathcal{S}, \xi) d\mathbf{x})^2\}$. Using Eq. (5.15) allows us to write $\sum_{i \leq j, \ell \leq r} (\int_{\Omega} \mathfrak{c}_{ij\ell r}(\mathbf{x}; \mathcal{S}, \xi) d\mathbf{x})^2 \leq |\Omega| \int_{\Omega} \|\mathfrak{c}(\mathbf{x}; \mathcal{S}, \xi)\|_F^2 d\mathbf{x} \leq |\Omega|^2 \gamma_{\mathfrak{C}}(\xi)^2$. Using Eq. (5.16) and (5.18) yields $\sum_{i \leq j, \ell \leq r} |\mathcal{L}^{ij}(\mathbf{u}^{\ell r}(\cdot; \mathcal{S}, \xi); \mathcal{S}, \xi)|^2 \leq 84 \underline{c}_\epsilon^{-2} |\Omega|^2 \gamma_{\mathfrak{C}}(\xi)^4$. It can then be deduced the inequality $\|\boldsymbol{\lambda}(\mathcal{S}, \xi)\|_2^2 \leq 168 \underline{c}_\epsilon^{-2} \gamma_{\mathfrak{C}}(\xi)^4 + 2 \gamma_{\mathfrak{C}}(\xi)^2$ and consequently, we have $E\{\|\tilde{\Lambda}\|_2^2\} = \int_{\mathcal{C}_{\mathcal{S}}} \int_{\mathcal{C}_{\xi}} \|\boldsymbol{\lambda}(\mathcal{S}, \xi)\|_2^2 P_{\mathcal{S}}(d\mathcal{S}) \otimes P_{\Xi}(d\xi) \leq 168 \underline{c}_\epsilon^{-2} E\{\Gamma_{\mathfrak{C}}^4\} + 2 E\{\Gamma_{\mathfrak{C}}^2\}$. Using Eq. (4.8) yields Eq. (5.19) with $\gamma_\lambda^2 = 168 \underline{c}_\epsilon^{-2} \underline{\gamma}_{4,\mathfrak{C}}^4 + 2 \underline{\gamma}_{2,\mathfrak{C}}^2$. \square

5.6. Brief comments about numerical aspects of stochastic solver. The Monte Carlo simulation method [23, 26] is used as stochastic solver. Let $\{(\mathcal{S}^\kappa, \xi^\kappa) \in \mathcal{C}_{\mathcal{S}} \times \mathcal{C}_{\xi}, \kappa = 1, \dots, \kappa_{\text{sim}}\}$ be κ_{sim} independent realizations of random variables (\mathcal{S}, Ξ) using the generator of probability measures $P_{\mathcal{S}}$ (see Section 5.5) and P_{Ξ} (see Section 5.4(i)). The spatial discretization of the weak formulation defined by Eq. (5.12) of the stochastic BVP and the discretization of Eq. (5.3) with Eq. (5.4) can be performed by the finite element method. Using Section 5.5, $\lambda_1(\mathcal{S}^\kappa, \xi^\kappa) \geq \dots \geq \lambda_6(\mathcal{S}^\kappa, \xi^\kappa) > 0$ are computed as the eigenvalues of $[\mathfrak{c}^{\text{eff}}(\mathcal{S}^\kappa, \xi^\kappa)] \in \mathbb{M}_6^+$. Taking into account Eq. (5.19), the mean-square convergence of the random eigenvalues can be analyzed with the convergence function

$$(5.20) \quad \kappa_{\text{sim}} \mapsto \text{conv}(\kappa_{\text{sim}}) = \|\underline{\mathfrak{C}}\|_F^{-1} \left(\kappa_{\text{sim}}^{-1} \sum_{\kappa=1}^{\kappa_{\text{sim}}} \|\boldsymbol{\lambda}(\mathcal{S}^\kappa, \xi^\kappa)\|_2^2 \right)^{1/2}.$$

For a given tolerance of convergence, the probability density function (pdf) $\tilde{\lambda} \mapsto p_{\tilde{\Lambda}}(\tilde{\lambda})$ on \mathbb{R}^6 (with support $(\mathbb{R}^{+*})^6$) with respect to the Lebesgue measure $d\tilde{\lambda}$ can be estimated with the κ_{sim} independent realizations $\{\boldsymbol{\lambda}(\mathcal{S}^\kappa, \xi^\kappa), \kappa = 1, \dots, \kappa_{\text{sim}}\}$ using, for instance, the multidimensional Gaussian kernel-density estimation method [3]. The pdf $\tilde{\lambda}_1 \mapsto p_{\tilde{\Lambda}_1}(\tilde{\lambda}_1)$ of $\tilde{\Lambda}_1$ can also be estimated yielding the pdf of the operator norm $\|\tilde{\mathfrak{C}}^{\text{eff}}\|_2 = \tilde{\Lambda}_1$.

6. NUMERICAL ILLUSTRATION

(i) *Mean model of the microstructure.* Domain $\Omega =]0, 1]^3$ and the mean model of the elastic material is chosen in the orthotropic class with mean Young moduli $\underline{E}_1 = 10^{10}$, $\underline{E}_2 = 0.5 \times 10^{10}$, and $\underline{E}_3 = 0.1 \times 10^{10}$, with mean Poisson coefficients $\underline{\nu}_{23} = 0.25$, $\underline{\nu}_{31} = 0.15$, and $\underline{\nu}_{12} = 0.1$ (the International System of Units is used).

(ii) *Elasticity random field.* The hyperparameter δ_c that allows for controlling the level of statistical fluctuations of $[\mathfrak{C}(\mathbf{x}; \mathcal{S})]$ (see Hypothesis 1) is fixed to the value 0.4.

(iii) *Uncertain spectral measure.* The model of the spectral measure is the one described in Example 3.5. The probability measure $P_{\mathbf{W}}(d\mathbf{w})$ of \mathbf{W} (see Definition 3.6) is chosen as uniform on \mathcal{C}_w . For all $j \in \{1, 2, 3\}$, the mean value of the random correlation length $L_{cj} = \pi/K_j$ is \underline{L}_c and its coefficient of variation $\delta_{L_{cj}}$ is δ_{L_c} , which are independent of j . Consequently, the support $[w_j^{\min}, w_j^{\max}]$ of the probability measure of $W_j = L_{cj}$ is such that $w_j^{\min} = \underline{L}_c (1 - \sqrt{3} \delta_{L_c})$ and $w_j^{\max} = 2 \underline{L}_c - w_j^{\min}$. The hyperparameter δ_s that controls the level of uncertainties of the spectral measure, which is such that $\delta_s^2 = (\prod_{j=1}^3 (1 + \delta_j^2)) - 1$,

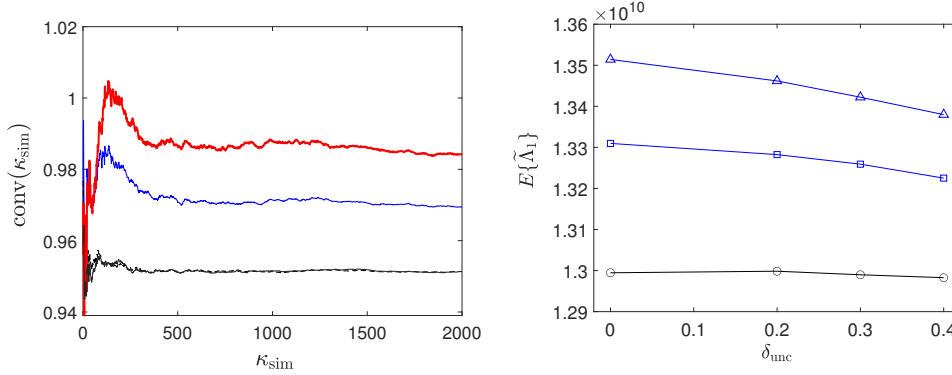


FIGURE 1. Left figure: function $\kappa_{\text{sim}} \mapsto \text{conv}(\kappa_{\text{sim}})$ for $\delta_{\text{unc}} = 0$ and $\underline{L}_c = 0.2$ (dashed line); $\delta_{\text{unc}} = 0.4$ and $\underline{L}_c = 0.2$ (black thin line), 0.4 (blue med line), and 0.6 (red thick line). Right figure: $E\{\tilde{\Lambda}_1\}$ as a function of δ_{unc} for $\underline{L}_c = 0.2$ (circle marker), 0.4 (square marker), and 0.6 (triangle-up marker).

is generated with $\delta_1 = \delta_2 = \delta_3 = \delta_{\text{unc}}$. A sensitivity analysis with respect to the level of spectrum uncertainties will be performed by considering 9 values of the triplet of parameters $(\underline{L}_c, \delta_{L_c}, \delta_{\text{unc}})$ with $\underline{L}_c \in \{0.2, 0.4, 0.6\}$ and $\delta_{L_c} = \delta_{\text{unc}} \in \{0.2, 0.3, 0.4\}$. For this set of data, the minimum of correlation lengths is 0.06 (obtained for $\underline{L}_c = 0.2$ with $\delta_{L_c} = \delta_{\text{unc}} = 0.4$) while the maximum is 1.01 (obtained for $\underline{L}_c = 0.6$ with $\delta_{L_c} = \delta_{\text{unc}} = 0.4$). The spectral domain sampling for the discretization of the spectral measure is performed with $\nu_s = 8$ and thus $\nu = 8^3 = 512$. With the quadrant symmetry, we have $\hat{\nu}_s = 4$ yielding $\hat{\nu} = 64$.

(iv) *Finite element discretization.* The weak formulation defined by Eq. (5.12) is discretized by the finite element method. The finite element mesh is made up of $20 \times 20 \times 20 = 8000$ solid finite elements (8-nodes solid), 9261 nodes, and 27783 degrees of freedom (dof). There are 2402 nodes on the boundary and thus 7206 zeros Dirichlet conditions. There are 2^3 integrations points in each finite element, which yields 64000 integrations points. The spatial discretization of the \mathbb{M}_6^+ -valued elasticity random field $[\mathbf{C}(\cdot, \mathcal{S})]$ yields $21 \times 64000 = 1344000$ random terms (taking into account the symmetry).

(v) *Stochastic solver.* The Monte Carlo simulation method is performed for $\kappa_{\text{sim}} \in [1, 2000]$. Figure 1-(left) displays the convergence function $\kappa_{\text{sim}} \mapsto \text{conv}(\kappa_{\text{sim}})$ defined by Eq. (5.20) for $\underline{L}_c = 0.2$ and with no uncertainty in the spectral measure ($\delta_{L_c} = \delta_{\text{unc}} = 0$) and for $\underline{L}_c \in \{0.2, 0.4, 0.6\}$ with the largest uncertainties in the spectral measure, $\delta_{L_c} = \delta_{\text{unc}} = 0.4$, which is the most unfavorable value with respect to convergence. It can be seen that the mean-square convergence is obtained for $\kappa_{\text{sim}} = 2000$.

(vi) *Sensitivity of the probability density function of the normalized operator norm of the random effective elasticity matrix as a function of the uncertainty level of the spectral measure.* Let $\Lambda_1 = \tilde{\Lambda}_1 / E\{\tilde{\Lambda}_1\}$ be the normalized operator norm $\|[\tilde{\mathbf{C}}^{\text{eff}}]\| / E\{[\tilde{\mathbf{C}}^{\text{eff}}]\| \} = \Lambda_1$. For $\underline{L}_c = 0.2, 0.3, 0.4$, Fig. 1-(right) displays the graph of $E\{\tilde{\Lambda}_1\}$ as a function of the level $\delta_{L_c} = \delta_{\text{unc}}$ of the spectral measure uncertainties. Figure 2 shows the pdf $\lambda_1 \mapsto p_{\Lambda_1}(\lambda_1)$ of the normalized operator norm of the random effective elasticity matrix for $\underline{L}_c = 0.2, 0.4$, and 0.6 , with no uncertainty in the spectral measure ($\delta_{L_c} = \delta_{\text{unc}} = 0$)

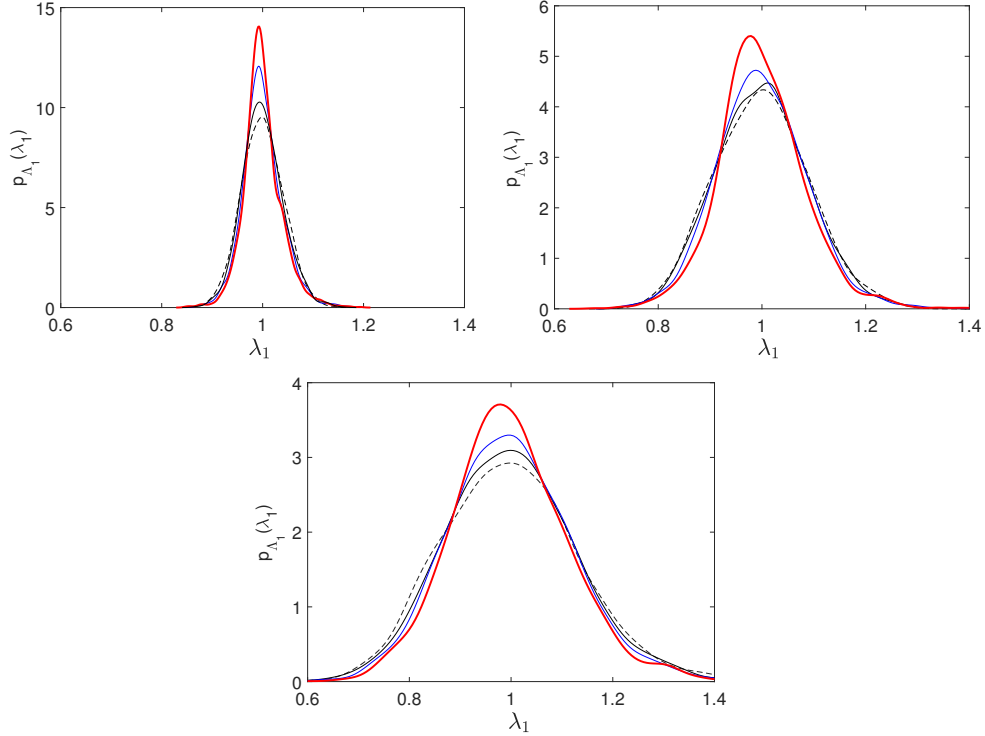


FIGURE 2. Sensitivity of the pdf $\lambda_1 \mapsto p_{\Lambda_1}(\lambda_1)$ for $\underline{L}_c = 0.2$ (top left figure), 0.4 (top right figure), and 0.6 (down figure) as a function of the spectral measure uncertainty level: $\delta_{\text{unc}} = 0$ (dashed line), 0.2 (black thin line), 0.3 (blue med line), 0.4 (red thick line).

and with uncertainties $\delta_{L_c} = \delta_{\text{unc}} = 0.2, 0.3,$ and 0.4 .

(vii) *Sensitivity of the probabilistic analysis of the RVE size with respect to the uncertainty level of the spectral measure.* We analyze the random largest eigenvalue Λ_1 (normalized operator norm). Let η be a positive real number and let $\eta \mapsto P(\eta)$ be the function from $]0, 1[$ into $[0, 1]$ defined by

$$(6.1) \quad P(\eta) = \text{Proba}\{1 - \eta < \Lambda_1 \leq 1 + \eta\} = F_{\Lambda_1}(1 + \eta) - F_{\Lambda_1}(1 - \eta),$$

in which F_{Λ_1} is the cumulative distribution function of Λ_1 . Figure 3 shows the sensitivity of the graph of function $\eta \mapsto P(\eta)$ for $\underline{L}_c = 0.2, 0.4,$ and 0.6 with respect to the level of uncertainties in the spectral measure for $\delta_{L_c} = \delta_{\text{unc}} = 0, 0.2, 0.3,$ and 0.4 . Table 1 yields an extraction from Fig. 3 of the probability levels.

(viii) *Brief discussion.* When there are no uncertainties in the spectral measure ($\delta_{\text{unc}} = 0$), Figure 3 and Table 1 shows that the condition to obtain a scale separation is not really obtained because, for $\underline{L}_c = 0.2$ and $\delta_{\text{unc}} = 0$ it can be seen that $\text{Proba}\{0.98 < \Lambda_1 \leq 1.02\} = 0.365$ and the probability becomes 0.942 only for $\{0.92 < \Lambda_1 \leq 1.08\}$. The results show in Table 1 shows that, for the specific case analyzed (in particular choosing the same level of uncertainties for the spatial correlation lengths and for the values of the spectral density function) and contrary to what was expected, the introduction of spectral measure uncertainties improves the scale separation from a probabilistic analysis point of

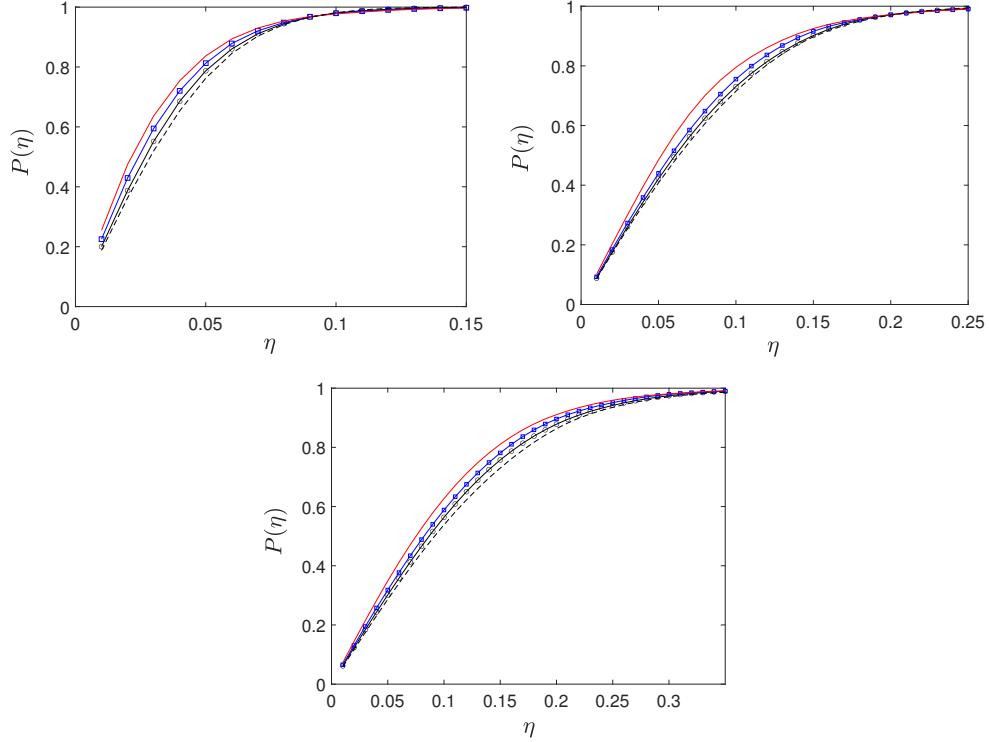


FIGURE 3. Sensitivity of the graph of function $\eta \mapsto P(\eta)$ for $\underline{L}_c = 0.2$ (top left figure), 0.4 (top right figure), and 0.6 (down figure) with respect to the level of uncertainties in the spectral measure for $\delta_{\text{unc}} = 0$ (dashed line), 0.2 (circle marker), 0.3 (square marker), and 0.4 (no marker).

TABLE 1. Sensitivity of the probabilistic analysis of the RVE size with respect to the uncertainty level of the spectral measure.

\underline{L}_c	δ_{unc}	$P\{0.98 < \Lambda_1 \leq 1.02\}$	$P\{0.96 < \Lambda_1 \leq 1.04\}$	$P\{0.92 < \Lambda_1 \leq 1.08\}$
0.2	0.0	0.365	0.655	0.942
	0.2	0.385	0.683	0.945
	0.3	0.430	0.721	0.948
	0.4	0.475	0.705	0.954
0.4	0.0	0.171	0.332	0.610
	0.2	0.175	0.344	0.625
	0.3	0.185	0.360	0.650
	0.4	0.202	0.395	0.700
0.6	0.0	0.108	0.230	0.442
	0.2	0.125	0.240	0.465
	0.3	0.130	0.260	0.488
	0.4	0.145	0.280	0.526

view. In fact, for $\underline{L}_c = 0.2$ and for $\delta_{L_c} = \delta_{\text{unc}} = 0.4$, the minimum value of the realizations of the random correlation length is 0.06, value less than 10 percent of the characteristic length of the dimensions of domain Ω , for which the scale separation can be obtained.

It should be noted that these results are presented as an illustration of the use of the proposed mathematical construction of a random field with uncertainties in the spectral measure in order to improve the probabilistic analysis of stochastic homogenization of random elastic media. More advanced computational analyses must be performed with this probabilistic model in order to deeply analyze the role played by uncertain spectral measure for stochastic homogenization of random elastic media. This work has been performed and can be found in [33].

BIBLIOGRAPHY

1. Robert Adler, *The geometry of random fields*, SIAM, 2010.
2. Michel Bornert, Thierry Bretheau, and Pierre Gilormini, *Homogenization in mechanics of materials*, ISTE Ltd and John Wiley and Sons, New York, 2008.
3. Adrian Bowman and Adelchi Azzalini, *Applied smoothing techniques for data analysis: The kernel approach with s-plus illustrations*, vol. 18, Oxford University Press, Oxford: Clarendon Press, New York, 1997.
4. Robert H Cameron and William T Martin, *The orthogonal development of non-linear functionals in series of Fourier-Hermite functionals*, Annals of Mathematics **48** (1947), no. 2, 385–392.
5. Robert Dautray and Jacques-Louis Lions, *Mathematical analysis and numerical methods for science and technology*, Springer Science & Business Media, Berlin, 2013.
6. Joseph Doob, *Stochastic processes*, John Wiley & Sons, New York, 1953.
7. Aleksandr D Egorov, P I Sobolevsky, and L A Yanovich, *Functional integrals: Approximate evaluation and applications*, Kluwer Academic Publishers, 1993.
8. Isosif Il'ich Guikhman and Anatoliĭ Skorokhod, *Introduction à la théorie des processus aléatoires*, Edition Mir, 1980.
9. Johann Guilleminot and Christian Soize, *On the statistical dependence for the components of random elasticity tensors exhibiting material symmetry properties*, Journal of Elasticity **111** (2013), no. 2, 109–130.
10. ———, *Stochastic model and generator for random fields with symmetry properties: application to the mesoscopic modeling of elastic random media*, Multiscale Modeling & Simulation (A SIAM Interdisciplinary Journal) **11** (2013), no. 3, 840–870.
11. Toufik Kanit, Samuel Forest, Isabelle Galliet, Valérie Mounoury, and Dominique Jeulin, *Determination of the size of the representative volume element for random composites: statistical and numerical approach*, International Journal of solids and structures **40** (2003), no. 13-14, 3647–3679.
12. Paul Krée and Christian Soize, *Mathematics of random phenomena*, Reidel Pub. Co, 1986, (first published by Bordas in 1983 and also published by Springer Science & Business Media in 2012).
13. Peter Lax and Arthur Milgram, *Parabolic equations: Contributions to the theory of partial differential equations*, Annals of Mathematical Studies (1954), no. 33, 167–190.
14. Nikolai Leonenko and Andriy Olenko, *Tauberian and abelian theorems for long-range dependent random fields*, Methodology and Computing in Applied Probability **15** (2013), no. 4, 715–742.
15. Jacques Louis Lions and Enrico Magenes, *Non-homogeneous boundary value problems and applications: Vol. 1*, vol. 181, Springer Science & Business Media, 2012.
16. Anatoliy Malyarenko and Martin Ostoja-Starzewski, *Statistically isotropic tensor random fields: correlation structures*, Mathematics and Mechanics of Complex Systems **2** (2014), no. 2, 209–231.
17. ———, *Tensor-valued random fields for continuum physics*, Cambridge University Press, 2018.
18. Anthony Nouy and Christian Soize, *Random field representations for stochastic elliptic boundary value problems and statistical inverse problems*, European Journal of Applied Mathematics **25** (2014), no. 3, 339–373.
19. Martin Ostoja-Starzewski, *Random field models of heterogeneous materials*, International Journal of Solids and Structures **35** (1998), no. 19, 2429–2455.
20. ———, *Material spatial randomness: From statistical to representative volume element*, Probabilistic engineering mechanics **21** (2006), no. 2, 112–132.
21. Martin Ostoja-Starzewski, X Du, ZF Khisaeva, and W Li, *Comparisons of the size of the representative volume element in elastic, plastic, thermoelastic, and permeable random microstructures*, International Journal for Multiscale Computational Engineering **5** (2007), no. 2, 73–82.

22. Fabrice Poirion and Christian Soize, *Numerical methods and mathematical aspects for simulation of homogeneous and non homogeneous gaussian vector fields*, Probabilistic Methods in Applied Physics (Paul Krée and Walter Wedig, eds.), Springer-Verlag, Berlin, 1995, pp. 17–53.
23. Christian Robert and George Casella, *Monte carlo statistical methods*, Springer Science & Business Media, 2005.
24. Murray Rosenblatt, *Stationary sequences and random fields*, Springer Science & Business Media, 2012.
25. Yu Rozanov, *Random fields and stochastic partial differential equations*, Kluwer Academic Publishers, 1998.
26. Reuven Rubinstein and Dirk Kroese, *Simulation and the Monte Carlo method*, Second Edition, John Wiley & Sons, New York, 2008.
27. Anatolii Volodimirovich Skorokhod and Mikhaïlo Iosipovich Yadrenko, *On absolute continuity of measures corresponding to homogeneous gaussian fields*, Theory of Probability & Its Applications **18** (1973), no. 1, 27–40.
28. Christian Soize, *A nonparametric model of random uncertainties on reduced matrix model in structural dynamics*, Probabilistic Engineering Mechanics **15** (2000), no. 3, 277–294.
29. ———, *Non gaussian positive-definite matrix-valued random fields for elliptic stochastic partial differential operators*, Computer Methods in Applied Mechanics and Engineering **195** (2006), no. 1-3, 26–64.
30. ———, *Tensor-valued random fields for meso-scale stochastic model of anisotropic elastic microstructure and probabilistic analysis of representative volume element size*, Probabilistic Engineering Mechanics **23** (2008), no. 2-3, 307–323.
31. ———, *Random vectors and random fields in high dimension: parametric model-based representation, identification from data, and inverse problems*, Handbook of Uncertainty Quantification (Roger Ghanem, David Higdon, and Houman Owhadi, eds.), vol. 2, Springer, Cham, Switzerland, 2017, pp. 883–936.
32. ———, *Uncertainty quantification*, Interdisciplinary Applied Mathematics, vol. 47, Springer, New York, 2017.
33. ———, *Computational stochastic homogenization of heterogeneous media from an elasticity random field having an uncertain spectral measure*, Computational Mechanics **published online** (2021), 1–19.
34. Erik Vanmarcke, *Random fields: Analysis and synthesis*, World Scientific, Singapore, 2010.
35. Mikhail Iosifovich Yadrenko, *Spectral theory of random fields*, Optimization Software, 1983.

MSME UMR 8208, UNIVERSITÉ GUSTAVE EIFFEL, 5 BD DESCARTES, 77454 MARNE-LA-VALLÉE,
FRANCE

Email address: christian.soize@univ-eiffel.fr

Received 01/September/2021