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## ERGODICITY OF THE FISHER INFINITESIMAL MODEL WITH QUADRATIC SELECTION

#### VINCENT CALVEZ, THOMAS LEPOUTRE, AND DAVID POYATO

ABSTRACT. We study the convergence towards a unique equilibrium distribution of the solutions to a time-discrete model with non-overlapping generations arising in quantitative genetics. The model describes the dynamics of a phenotypic distribution with respect to a multi-dimensional trait, which is shaped by selection and Fisher's infinitesimal model of sexual reproduction. We extend some previous works devoted to the time-continuous analogues, that followed a perturbative approach in the regime of weak selection, by exploiting the contractivity of the infinitesimal model operator in the Wasserstein metric. Here, we tackle the case of quadratic selection by a global approach. We establish uniqueness of the equilibrium distribution and exponential convergence of the renormalized profile. Our technique relies on an accurate control of the propagation of information across the large binary trees of ancestors (the pedigree chart), and reveals an ergodicity property, meaning that the shape of the initial datum is quickly forgotten across generations. We combine this information with appropriate estimates for the emergence of Gaussian tails and propagation of quadratic and exponential moments to derive quantitative convergence rates. Our result can be interpreted as a generalization of the Krein-Rutman theorem in a genuinely non-linear, and non-monotone setting.

#### 1. Introduction

Fisher's infinitesimal model (also known as the polygenic model) is a widely used statistical model in quantitative genetics initially proposed by R. FISHER [13]. It assumes that the genetic component of a quantitative phenotypical trait is affected by an infinite number of loci with infinitesimal and additive allelic effects and claims that the genetic component of descendants' traits is normally distributed around the mean value of parents' traits, with a constant (genetic) variance across generations. This model allowed reconcilling Mendelian inheritance and the continuous trait variations documented by F. Galton via a Central Limit Theorem. More specifically, by taking limits when the number of underlying loci tends to infinity on a model with Mendelian inheritance, N. Barton, A. Etheridge and A. Véber [4] recently proved rigorously the validity of Fisher's infinitesimal model under various evolutionary processes (e.g., natural selection).

In this paper we study a time-discrete evolution problem for the distribution of a phenotypical trait  $x \in \mathbb{R}^d$  in a population undergoing sexual reproduction and the effect of natural selection. Specifically, starting at any initial configuration  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  of trait distribution, we analyze the long term dynamics of trait distributions  $\{F_n\}_{n\in\mathbb{N}}$  across successive generations  $n \in \mathbb{N}$ , which solves the following recursion

$$F_n = \mathcal{T}[F_{n-1}],\tag{1.1}$$

for any  $n \in \mathbb{N}$ . The operator  $\mathcal{T}$  encodes the balanced effect of sexual reproduction in the population (under the infinitesimal model) and natural selection. Specifically,  $\mathcal{T}$  is defined by

$$\mathcal{T}[F] := e^{-m} \mathcal{B}[F], \tag{1.2}$$

for any trait distribution  $F \in \mathcal{M}_+(\mathbb{R}^d)$ . On the one hand,  $m = m(x) \geq 0$  is called the selection function and represents the mortality effect of a trait-dependent natural selection on the population, so that  $e^{-m}$  stands for the survival probability in the next generation. On the other hand, the operator  $\mathcal{B}$  is chosen to be Fisher's infinitesimal operator and it takes the form

$$\mathcal{B}[F](x) := \int_{\mathbb{R}^{2d}} G\left(x - \frac{x_1 + x_2}{2}\right) \frac{F(x_1)F(x_2)}{\int_{\mathbb{R}^d} F(x') \, dx'} \, dx_1 \, dx_2, \quad x \in \mathbb{R}^d, \tag{1.3}$$

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for any trait distribution  $F \in \mathcal{M}_+(\mathbb{R}^d)$ . Here G = G(x) is a probability density (the *mixing kernel*) and the factor  $G(x - \frac{x_1 + x_2}{2})$  represents the transition probability that two given individuals with trait values  $x_1, x_2 \in \mathbb{R}^d$  will mate and yield a descendant with trait value  $x \in \mathbb{R}^d$ . In other words the resulting trait distributes around the mean value  $\frac{x_1 + x_2}{2}$  of parents' trait with law G. By definition,  $\mathcal{B}[F]$  quantifies the number of births after all possible matches of any couple of individuals according to the trait distribution F. Altogether,  $\mathcal{T}[F]$  quantifies the amount of offspring of a population distributed according to F having resisted the effect of selection.

The above sexual reproduction operator  $\mathcal{B}$  has recently pulled the attention of both the applied and more theoretical communities, cf. [4, 9, 15, 16, 17, 18]. In this paper we shall restrict to Gaussian mixing kernel and quadratic selection function, *i.e.*,

$$G(x) := \frac{1}{(2\pi)^{d/2}} e^{-\frac{|x|^2}{2}}, \quad x \in \mathbb{R}^d,$$
 (1.4)

$$m(x) := \frac{\alpha}{2}|x|^2, \qquad x \in \mathbb{R}^d, \tag{1.5}$$

where  $\alpha \in \mathbb{R}_+$  is a fixed parameter. Thereby, the trait of offsprings is normally distributed around the mean value of the trait of parents by assumption (1.4), thus reducing to the standard infinitesimal model when the assumptions of the Central Limit Theorem are met [4]. For simplicity, we set a Gaussian G with unit variance, but any value of the genetic variance could also be considered (see nondimensionalization in Appendix A). In addition, note that by assumption (1.5) the mortality rate is minimized at x = 0 and it grows quadratically as  $|x| \to \infty$ . In this study, we restrict to trait-dependent mortality, and consider a constant fecundity rate, for the sake of simplicity. We believe that the techniques developed here can easily be extended to a trait-dependent fecundity, provided that selection effects have a Gaussian shape, as for the survival probability  $e^{-m}$ .

Before introducing our results, we shall relate the previous time-discrete problem (1.1) to analogous time-continuous quantitative genetics models of evolutionary dynamics that have been studied in the literature. Meanwhile, we will anticipate the major difficulties that can be faced when analyzing the long-time dynamics of (1.1). To this end, we consider, the following type of integro-differential equations for the evolutionary dynamics of a trait distribution f(t, x):

$$\begin{cases}
\partial_t f = -m(x)f + \mathcal{R}[f], & t \ge 0, x \in \mathbb{R}^d, \\
f(0, x) = f_0(x), & x \in \mathbb{R}^d,
\end{cases}$$
(1.6)

where, m = m(x) is again the mortality rate and  $\mathcal{R} = \mathcal{R}[f]$  is the reproduction operator. Hence,  $\mathcal{R}[f](x)$  determines the amount of births with trait value  $x \in \mathbb{R}^d$  per unit of time. As for (1.1), the resulting dynamics of the population becomes a consequence of a balance between selection encoded by the trait-dependent mortality and the diversity generated by the growth term  $\mathcal{R}$  across generations.

Many studies consider a linear reproduction operator  $\mathcal{R}$ , associated with a probability density K = K(x) characterizing the mutational effects at birth, of the form

$$\mathcal{R}[f](x) := \int_{\mathbb{R}^d} K(x - y) f(t, y) \, dy, \quad x \in \mathbb{R}^d.$$
 (1.7)

The factor K(x-y) determines the probability that an individual with trait value  $y \in \mathbb{R}^d$  produces a descendant with trait value  $x \in \mathbb{R}^d$  (possibly deviating from y). This class of linear reproduction operators is well-suited for an asexual mode of reproduction. This includes parabolic equations in the limit of small variance of K. In particular, we refer to a series of works about the long-time asymptotics in the regime of small variance, initiated in [2, 3, 12], including an additional density-dependent competition term that makes the analysis non standard (see e.g. [10] and references therein for the well-posedness of the constrained Hamilton-Jacobi equation derived in the limit).

Recently, inspired by the infinitesimal model, the case of sexual reproduction has been addressed by invoking the preceding nonlinear version  $\mathcal{B}$  in (1.3) as reproduction operator  $\mathcal{R} = \mathcal{B}$ . Several asymptotic regimes have been addressed: large reproduction rate [15, 17, 18], small variance asymptotics [9, 16]. In the latter case, the limiting problem keeps the non-local nature of the problem, being of a finite-difference type, rather than a Hamilton-Jacobi PDE. Before we continue the discussion about the state-of-the-art, let us emphasize that there is no restriction on the parameters in the present work.

Since both  $f \mapsto mf$  and  $f \mapsto \mathcal{B}[f]$  are 1-homogeneous operators, we can seek for special steady solutions of (1.3) through the following ansatz:

$$f(t,x) = e^{\lambda t} F(x), \quad (t,x) \in \mathbb{R}_+ \times \mathbb{R}^d,$$
 (1.8)

The parameter  $\lambda \in \mathbb{R}$  represents the rate at which the number of individuals grows (if  $\lambda \geq 0$ ) or decreases (if  $\lambda \leq 0$ ), and  $F = F(x) \geq 0$  is an unknown probability density. By imposing such an ansatz on (1.6)-(1.3), the following generalized spectral problem arises for the couple  $(\lambda, F)$ :

$$\begin{cases} \lambda F(x) + m(x)F(x) = \mathcal{B}[F](x), & x \in \mathbb{R}^d, \\ \int_{\mathbb{R}^d} F(x') dx' = 1. \end{cases}$$
 (1.9)

Note that the operator  $\mathcal{B}$  is genuinely non-linear so that methods based on the Krein-Rutman theory or maximum principles cannot be applied straightforwardly, see [5, 11] and the references therein for the linear case. Further, usual extensions of the Krein-Rutman theory to 1-homogeneous operators [14] cannot be applied neither because  $\mathcal{B}$  is not monotone. To date, the main strategy behind the existence of solutions of (1.9) relies on a suitable application of Schauder fixed-point theorem to the operator  $F \mapsto (\lambda + m)^{-1}\mathcal{B}[F]$  over an appropriate cone of  $L^1(\mathbb{R}^d)$  that is conserved by the nonlinear operator, see [7]. However, uniqueness cannot be achieved by this method. Moreover, it has been proven in [9, Corollary 1.5] that several equilibrium states  $(\lambda, F)$  can co-exist in the presence of multiple local minima of m (provided the variance of G is small enough). That is, the generalized eigenproblem (1.9) does not admit a unique positive eigenfunction, in contrast with general conclusions of the Krein-Rutman theory.

Recently, G. RAOUL addressed the long-time dynamics of (1.6) in 1D, with  $\mathcal{R} = \mathcal{B}$  and a trait dependent fecundity rate [18]. He obtained local uniqueness and exponential relaxation under the assumption of weak and localized (compactly supported) selection effects. To this end, he controlled locally in space the Wasserstein distance between the solution and the stationary state, using the uniform contraction property of  $\mathcal{B}$  in the space of probability measures sharing the same center of mass. Unfortunately, the Wasserstein metric is not fully compatible with multiplicative operators, such as trait-dependent fecundity (see the discussion in [18, Section 3.4] and Section 2 below). This can be circumvented under the additional assumption that the trait density is locally uniformly bounded below, following [6]. Obviously, this cannot hold globally for integrable densities, hence G. RAOUL developed estimates of the distribution's tail to complete the contraction estimates. Also, a lot of attention has to be paid to the dynamics of the center of the distribution which is essentially driven by selection. Indeed, in the case of flat selection  $(m \equiv 0)$ , the problem is invariant by translation, so that local uniqueness cannot hold. To conclude this discussion, let us emphasize that global uniqueness and the asymptotic behavior of generic solutions to the evolution problem (1.6)-(1.3) is still open. Below, we provide a first result in this direction, for the time-discrete problem (1.1), though.

We remark that the time-discrete version (1.1) that we propose in this paper can be partially regarded as a discretization in time of the above time-continuous problem (1.6)-(1.3), with non-overlapping generations (see Appendix A for further details). As for the time-continuous problem, we could seek special solutions to the time-discrete problem (1.1) in the following form

$$F_n(x) = \lambda^n F(x), \quad (n, x) \in \mathbb{N} \times \mathbb{R}^d.$$
 (1.10)

Again,  $\lambda \in \mathbb{R}_+^*$  is the rate of growth (if  $\lambda \geq 1$ ) or decrease (if  $\lambda \leq 1$ ) of individuals, and F = F(x) is an unknown probability density. This yields the following generalized eigenproblem for the couple  $(\lambda, F)$ :

$$\begin{cases} \lambda F(x) = \mathcal{T}[F](x), & x \in \mathbb{R}^d, \\ \int_{\mathbb{R}^d} F(x') \, dx' = 1. \end{cases}$$
 (1.11)

In this paper we aim to address the following questions:

- (Q1) Does the eigenproblem (1.11) have a unique solution  $(\lambda_{\alpha}, \mathbf{F}_{\alpha})$  for each  $\alpha \in \mathbb{R}+?$
- (Q2) Consider any generic initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  and its associated solution  $\{F_n\}_{n\in\mathbb{N}}$  of the time-discrete problem (1.1). Do the renormalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  converge to the unique steady profile  $F_{\alpha}$  solving (1.11) when  $n \to \infty$ ?

We shall prove that the answer to both questions is affirmative. It stands to reason that similar existence and local uniqueness results like in [7, 9] could be extended to the new eigenproblem (1.11) by applying Schauder and Banach fixed point theorems. Nevertheless, in this paper we introduce a novel method that unravels an ergodicity property of the operator  $\mathcal{T}$ , leading to quantitative estimates for the relaxation of profiles  $\{F_n\}_{n\in\mathbb{N}}$  towards  $\mathbf{F}_{\alpha}$ . First, we prove that a explicit (Gaussian) solution to the eigenproblem (1.11) exists. Second, by computing n iterations of the operator  $\mathcal{T}$ , that is  $F_n = \mathcal{T}^n[F_0]$ , we notice that information of  $F_n$  at the trait value x is propagated from the initial datum  $F_0$  across  $2^n$  ancestors over a binary tree with height n and rooted at x (the pedigree chart). Interestingly, an appropriate reformulation of  $\mathcal{T}^n$  in the case of Gaussian mixing (1.4) and quadratic selection (1.5) shows that the dependence of the solution  $\{F_n\}_{n\in\mathbb{N}}$  on the initial datum  $F_0$  is rapidly lost across the different levels of the tree. More specifically, a strong convergence of generic solutions  $\{F_n\}$  of the time-discrete problem (1.1) towards

the steady Gaussian profile  $F_{\alpha}$  solving (1.11) is achieved locally with respect to x. Third, we prove an appropriate propagation of quadratic and exponential moments, leading to uniform tightness of the family  $\{F_n\}_{n\in\mathbb{N}}$ . Finally, we glue all the information together and conclude the final global convergence result in question (Q2) in relative entropy. Specifically, we obtain our main result:

**Theorem 1.1.** Assume that  $\alpha \in \mathbb{R}_+^*$  and set any initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$ . Hence, the solution  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1) verifies that the growth rates  $\|F_n\|_{L^1(\mathbb{R}^d)}/\|F_{n-1}\|_{L^1(\mathbb{R}^d)}$  relax towards  $\lambda_{\alpha}$  and the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  relax towards  $F_{\alpha}$ , where  $(\lambda_{\alpha}, F_{\alpha})$  are given by

$$\lambda_{\alpha} := \left(1 + \alpha \left(1 + \frac{\sigma_{\alpha}^{2}}{2}\right)\right)^{-\frac{d}{2}}, \quad \boldsymbol{F}_{\alpha} := G_{0,\sigma_{\alpha}^{2}}, \tag{1.12}$$

and the variance  $\sigma_{\alpha}^2 \in \mathbb{R}_+^*$  is the unique positive root of the equation

$$\frac{1}{\sigma_{\alpha}^2} = \alpha + \frac{1}{1 + \frac{\sigma_{\alpha}^2}{2}}, \quad i.e., \quad \sigma_{\alpha}^2 = \frac{\sqrt{(1 + 2\alpha)^2 + 8\alpha} - (1 + 2\alpha)}{2\alpha}.$$
 (1.13)

Specifically, for any  $\varepsilon \in \mathbb{R}_+^*$  there exists a sufficiently large  $C_\varepsilon \in \mathbb{R}_+^*$  such that

$$\left| \frac{\|F_n\|_{L^1(\mathbb{R}^d)}}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} - \boldsymbol{\lambda}_{\alpha} \right| \le C_{\varepsilon} ((2\,\boldsymbol{k}_{\alpha})^{1/3} + \varepsilon)^n,$$

$$\mathcal{D}_{KL} \left( \frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} \right\| \boldsymbol{F}_{\alpha} \right) \le C_{\varepsilon} ((2\,\boldsymbol{k}_{\alpha})^{2/3} + \varepsilon)^n,$$

for any  $n \in \mathbb{N}$ , where  $\mathcal{D}_{KL}$  is the Kullback-Leibler divergence (or relative entropy) and  $\mathbf{k}_{\alpha} \in (0, \frac{1}{2})$  reads

$$\mathbf{k}_{\alpha} = \frac{(3+2\alpha) - \sqrt{(3+2\alpha)^2 - 8}}{4}.$$
(1.14)

Note that any solution  $(\lambda, F)$  to the eigenproblem (1.11) yields a solution to the time-discrete problem (1.1) through the ansatz (1.10). Then, question (Q1) will readily follow from question (Q2) and, in particular, the unique solution  $(\lambda_{\alpha}, F_{\alpha})$  to (1.11) is Gaussian.

Corollary 1.2. Assume that  $\alpha \in \mathbb{R}_+^*$ , then there exists a unique solution to the eigenproblem (1.11) and it is given by  $(\lambda_{\alpha}, \mathbf{F}_{\alpha})$  where  $\lambda_{\alpha}$  and  $\mathbf{F}_{\alpha}$  are determined by (1.12).

On the one hand, note the necessity of  $\alpha \in \mathbb{R}_+^*$  in both Theorem 1.1 and Corollary 1.2. Indeed, for  $\alpha = 0$  we recover infinitely many solutions by the translation invariance of the infinitesimal operator  $\mathcal{B}$  in (1.3). On the other hand, we emphasize that our choice of Gaussian mixing kernel (1.4) and the quadratic selection function (1.5) will be used throughout the paper and it will allow for some explicit calculations. In particular, the unique solution of the eigenproblem (1.11) is the explicit Gaussian profile determined by  $(\lambda_{\alpha}, F_{\alpha})$  in (1.12). However, we believe that the Gaussian structure is not mandatory and a similar ergodicity property might be true for more general mixing kernels G and selection functions m with a less suited algebraic structure. We leave all these extensions, along with the case of overlapping generations and the full time-continuous dynamics, to future works.

The rest of the paper is organized as follows. In Section 2 we discuss about the generic incompatibility of the Wasserstein distance with a multiplicative operator and we provide a brief outlook of the strategy of our proof. In Section 3 we provide some necessary notation and we introduce the special class of Gaussian solutions of both problems (1.1) and (1.11), which will inspire some parts of the paper. Section 4 is devoted to introduce some main properties of  $\mathcal{T}$  regarding the emergence of Gaussian behavior (in the large) from generic initial data, and a suitable propagation of quadratic and exponential moments across generations. In Section 5 we reformulate (1.1) via a high-dimensional integral operator propagating ancestors' information across the different levels of the pedigree chart, which will be the cornerstone to study the long-term dynamics. In Section 6 we prove our main results, namely, Theorem 1.1 and Corollary 1.2. Section 7 contains some numerical experiments that illustrate the results in this paper. In Section 8 we provide some conclusions and perspectives. Finally, Appendix A contains the adimensionalization of the problem, and the relationship between (1.1) and the previous time-continuous analogues in the literature is discussed.

#### Notation:

• Euclidean norm: We denote the Euclidean norm and Euclidean inner product by |x| and  $x \cdot y$  respectively for any  $x, y \in \mathbb{R}^d$ . Similarly, we denote  $\|\mathbf{x}\| = (\sum_{i=1}^n |x_i|^2)^{1/2}$  for any  $\mathbf{x} = (x_1, \dots, x_n) \in \mathbb{R}^{nd}$  with  $x_i \in \mathbb{R}^d$ .

- Function spaces:  $\mathcal{M}(\mathbb{R}^d)$  denotes the Banach space of finite Radon measures on  $\mathbb{R}^d$  endowed with the total variation norm and  $\mathcal{M}_+(\mathbb{R}^d)$  represents the subspace of non-negative Radon measures. Similarly,  $\mathcal{P}(\mathbb{R}^d)$  is the metric space consisting of probability measures endowed with narrow topology, and  $\mathcal{P}_2(\mathbb{R}^d)$  is the subset of probability measures with finite second order moment endowed with the quadratic Wasserstein distance  $\mathcal{W}_2$ . Finally, for any  $p \in [1, \infty]$  and  $k \in \mathbb{N}$ ,  $L^p(\mathbb{R}^d)$  and  $W^{k,p}(\mathbb{R}^d)$  denote the usual Lebesgue and Sobolev spaces.
- Expectations: Given any random variable  $X:(\Omega,\mathcal{F},\mathbb{P})\longrightarrow\mathbb{R}$ , we denote its expectation by  $\mathbb{E}[X]$ .
- Gaussian functions: Given  $\mu \in \mathbb{R}^d$  and a positive definite symmetric matrix  $\Sigma \in \mathcal{M}_d(\mathbb{R}^d)$ , we will denote the density of the multivariate normal distribution with mean  $\mu$  and covariance matrix  $\Sigma$  by

$$G_{\mu,\Sigma}(x) := \frac{1}{\sqrt{(2\pi)^d \det(\Sigma)}} e^{-\frac{1}{2}(x-\mu)^\top \Sigma^{-1}(x-\mu)}, \quad x \in \mathbb{R}^d.$$

As usual, a random vector  $X:(\Omega,\mathcal{F},\mathbb{P})\longrightarrow\mathbb{R}^d$  will be said normally distributed with mean  $\mu$  and covariance matrix  $\Sigma$  (in symbols  $X\sim\mathcal{N}(\mu,\Sigma)$ ) if  $\mathrm{Law}(X)=G_{\mu,\Sigma}$ . When  $\Sigma=\sigma^2I_d$  for some  $\sigma\in\mathbb{R}_+^*$ , we will simplify notation  $G_{\mu,\sigma^2}\equiv G_{\mu,\sigma^2I_d}$  and it will be called the Gaussian centered at  $\mu$  with variance  $\sigma^2$ .

#### 2. Motivation and strategy of the proof of Theorem 1.1

In this section, we discuss the incompatibility of the quadratic Wasserstein distance to quantify directly contractivity under the joint effect of the reproduction operator  $\mathcal{B}$  in (1.3) and a generic multiplicative selection  $e^{-m}$ . Although a small perturbation of the case of flat selection  $(m \equiv 0)$  could still be considered via a perturbative argument (see discussion above, [18] and also [17]), our novel approach is able to tackle a purely non-perturbative setting. We end the section by briefly discussing the strategy of our proof.

2.1. Some properties of the sexual reproduction operator. We start by recalling some of the main properties of the sexual reproduction operator  $\mathcal{B}$ . Since the fecundity rate has been normalized to 1 (see Appendix A), then  $\mathcal{B}$  preserves the mass and center of mass, namely,

$$\|\mathcal{B}[F]\|_{L^1(\mathbb{R}^d)} = \|F\|_{\mathcal{M}(\mathbb{R}^d)}, \quad \int_{\mathbb{R}^d} x \, \mathcal{B}[F](x) \, dx = \int_{\mathbb{R}^d} x \, F(dx),$$
 (2.1)

for any  $F \in \mathcal{M}_+(\mathbb{R}^d)$ . Furthermore, it is contractive in the space of probability measures with a common center of mass, endowed with the quadratic Wasserstein metric. As discussed above, this property has been used fruitfully by G. RAOUL (cf. [18]) to analyse the long term behavior of the time-continuous problem in the regime of weak (and compactly supported) selection acting on fecundity. For the sake of clarity, we recall this fact and its proof below:

**Lemma 2.1.** Assume that  $F_1, F_2 \in \mathcal{P}_2(\mathbb{R}^d)$  and they have the same center of mass. Then,

$$W_2^2(\mathcal{B}[F_1], \mathcal{B}[F_2]) \le \frac{1}{2} W_2^2(F_1, F_2).$$

Proof. Recall the following dual characterizations of the quadratic Wasserstein distance (cf. [1, 21]):

$$\mathcal{W}_{2}^{2}(F_{1}, F_{2}) = \inf \left\{ \int_{\mathbb{R}^{2d}} |x - y|^{2} \gamma(dx, dy) : \gamma \in \mathcal{P}(\mathbb{R}^{d} \times \mathbb{R}^{d}), \ \pi_{1\#} \gamma = F_{1}, \ \pi_{2\#} \gamma = F_{2} \right\} 
= \sup \left\{ \int_{\mathbb{R}^{d}} \phi_{1} dF_{1} + \int_{\mathbb{R}^{d}} \phi_{2} dF_{2} : |\phi_{1}(x) + \phi_{2}(y)| \leq |x - y|^{2} \ \forall x, y \in \mathbb{R}^{d} \right\},$$
(2.2)

where  $\pi_i : \mathbb{R}^d \times \mathbb{R}^d \to \mathbb{R}^d$  is the projection onto the *i*-th component for i = 1, 2. Taking any couple  $(\phi_1, \phi_2)$  as above and using the specific form of the operator  $\mathcal{B}$  we obtain:

$$\begin{split} & \int_{\mathbb{R}^d} \mathcal{B}[F_1](x) \,\phi_1(x) \,dx + \int_{\mathbb{R}^d} \mathcal{B}[F_2](y) \,\phi_2(y) \,dy \\ & = \int_{\mathbb{R}^{3d}} G\left(x - \frac{x_1 + x_2}{2}\right) \,\phi_1(x) \,F_1(dx_1) \,F_1(dx_2) \,dx + \int_{\mathbb{R}^{3d}} G\left(y - \frac{y_1 + y_2}{2}\right) \,\phi_2(y) \,F_2(dy_1) \,F_2(dy_2) \,dy \\ & = \int_{\mathbb{R}^{3d}} G(z) \,\phi_1\left(z + \frac{x_1 + x_2}{2}\right) \,F_1(dx_1) \,F_1(dx_2) \,dz + \int_{\mathbb{R}^{3d}} G(z) \,\phi_2\left(z + \frac{y_1 + y_2}{2}\right) \,F_2(dy_1) \,F_2(dy_2) \,dz. \end{split}$$

Consider any transference plan  $\gamma \in \mathcal{P}(\mathbb{R}^d \times \mathbb{R}^d)$  between  $(F_1, F_2)$  as above. Specifically, we have that  $\pi_{1\#}\gamma = F_1$  and  $\pi_{2\#}\gamma = F_2$ . Then, we can gather both integrals as follows:

$$\int_{\mathbb{R}^d} \mathcal{B}[F_1](x) \,\phi_1(x) \,dx + \int_{\mathbb{R}^d} \mathcal{B}[F_2](y) \,\phi_2(y) \,dy 
= \int_{\mathbb{R}^{5d}} G(z) \left( \phi_1 \left( z + \frac{x_1 + x_2}{2} \right) + \phi_2 \left( z + \frac{y_1 + y_2}{2} \right) \right) \,\gamma(dx_1, dy_1) \,\gamma(dx_2, dy_2) \,dz.$$

Since the condition  $|\phi_1(x) + \phi_2(y)| \leq |x - y|^2$  is verified for all  $x, y \in \mathbb{R}^d$ , then we find

$$\int_{\mathbb{R}^{d}} \mathcal{B}[F_{1}](x) \,\phi_{1}(x) \,dx + \int_{\mathbb{R}^{d}} \mathcal{B}[F_{2}](y) \,\phi_{2}(y) \,dy$$

$$\leq \frac{1}{4} \int_{\mathbb{R}^{5d}} G(z) \left| (x_{1} + x_{2}) - (y_{1} + y_{2}) \right|^{2} \gamma(dx_{1}, dy_{1}) \gamma(dx_{2}, dy_{2}) \,dz$$

$$= \frac{1}{4} \int_{\mathbb{R}^{4d}} \left| (x_{1} + x_{2}) - (y_{1} + y_{2}) \right|^{2} \gamma(dx_{1}, dy_{1}) \gamma(dx_{2}, dy_{2})$$

$$= \frac{1}{4} \int_{\mathbb{R}^{2d}} \left| x_{1} - y_{1} \right|^{2} \gamma(dx_{1}, dy_{1}) + \frac{1}{4} \int_{\mathbb{R}^{2d}} \left| x_{2} - y_{2} \right|^{2} \gamma(dx_{2}, dy_{2}),$$

where in the third line we have used that G is a probability density and in the last line we have used the crucial fact that  $F_1$  and  $F_2$  share the same center of mass in order to cancel the cross-terms (otherwise the estimate would boil down to a non-expansiveness estimate). Taking supremum over  $(\phi_1, \phi_2)$  and infimum over  $\gamma$  and using the dual characterizations (2.2) yields the result.

Note that the fact that both  $F_1$  and  $F_2$  have the same center of mass has been crucially used to cancel the crossed term. Otherwise, by the Cauchy–Schwarz inequality we would merely obtain non-expansiveness:

$$\mathcal{W}_2(\mathcal{B}[F_1], \mathcal{B}[F_2]) \le \mathcal{W}_2(F_1, F_2). \tag{2.3}$$

Using the contractivity property of Lemma 2.1 along with the conservation of mass and center of mass in (2.1) yields the long term dynamics of (1.1) in the special case  $\alpha = 0$ .

Corollary 2.2. Set any initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  such that  $\int_{\mathbb{R}^d} |x|^2 F_0(dx) < \infty$  and consider the solution  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1) with  $\alpha=0$ , i.e.,  $F_n=\mathcal{B}[F_{n-1}]$  for all  $n\in\mathbb{N}$ . Then,

$$\frac{\|F_n\|_{L^1(\mathbb{R}^d)}}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} = 1, \quad \mathcal{W}_2\left(\frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}}, G_{\mu_0, 2}\right) \lesssim \frac{1}{2^{n/2}},$$

for every  $n \in \mathbb{N}$ , where  $\mu_0 := \int_{\mathbb{R}^d} x \, dF_0(x)$ . In particular, the set of stationary distributions under  $\mathcal{B}$  is  $\{F_0(\cdot - \mu) : \mu \in \mathbb{R}^d\}$ , where  $F_0 = G_{0,2}$  will denote here on the Gaussian centered at the origin with variance equals 2 in agreement with the notation (1.12) in Theorem 1.1.

The preceding result might be regarded as the counterpart of Theorem 1.1 for  $\alpha = 0$ . Indeed, by Talagrand transportation inequality for a Gaussian measure [20] we obtain the relation

$$\mathcal{W}_{2}^{2}\left(\frac{F_{n}}{\|F_{n}\|_{L^{1}(\mathbb{R}^{d})}}, G_{\mu_{0}, 2}\right) \leq 4 \, \mathcal{D}_{\mathrm{KL}}\left(\frac{F_{n}}{\|F_{n}\|_{L^{1}(\mathbb{R}^{d})}} \left\| G_{\mu_{0}, 2}\right).$$

However, we remark that Theorem 1.1 does not hold if  $\alpha=0$  due to two fundamental reasons. On the one hand, note that by the previous Corollary 2.2,  $F_n$  converge towards a Gaussian with the same center of mass as the initial datum. Then, the centers of mass of  $F_n$  do not necessarily get shifted towards the origin in this case. This is in agreement with the translation invariance of  $\mathcal{B}$ . In particular, that explains why equilibria are not unique for  $\alpha=0$  contrarily to the case  $\alpha\in\mathbb{R}_+^*$ . On the other hand, even if we set the initial center of mass at the origin (i.e.,  $\mu_0=0$ ) so that we kill the translation invariance, our method of proof of Theorem 1.1 leads to estimates that blow up as  $\alpha\to 0$ .

2.2. Incompatibility of Wasserstein metric with multiplicative operators. Note that by definition (1.2), our operator  $\mathcal{T}$  is the composition of the sexual resproduction operator  $\mathcal{B}$  with the multiplicative operator by the survival probability  $e^{-m}$ . Then, it might be natural to study perturbations of the previous Lemma 2.1 including the following conservative multiplicative operator.

**Definition 2.3** (Normalization of multiplicative operator).

$$\mathcal{M}[F] := \frac{e^{-m}F}{\|e^{-m}F\|_{\mathcal{M}(\mathbb{R}^d)}}, \quad F \in \mathcal{M}_+(\mathbb{R}^d) \setminus \{0\}.$$

However, the latter is not Lipschitz continuous with respect to the quadratic Wasserstein metric. Hence the composition of  $\mathcal{B}$  and  $\mathcal{M}$  is not expected to be contractive, even in the case of weak selection, without any additional restriction. We illustrate such a Lipschitz discontinuity of  $\mathcal{M}$  in the following example.

**Example 2.4.** Suppose that the  $m \in C^1_+(\mathbb{R})$  is radially symmetric around the origin and set  $F_1, F_2 \in \mathcal{P}_2(\mathbb{R})$  as the sum of two Dirac masses,  $F_1$  being symmetric, and  $F_2$  nearly symmetric. More precisely,

$$F_1 = \frac{1}{2}\delta_{-h} + \frac{1}{2}\delta_h, \quad F_2 = \frac{1}{2}\delta_{-h+\varepsilon} + \frac{1}{2}\delta_{h+\varepsilon},$$

where h > 0 is fixed so that  $m'(h) \neq 0$  and  $\varepsilon > 0$  is small. Note that

$$\mathcal{M}[F_1] = F_1, \quad \mathcal{M}[F_2] = (1 - p_{\varepsilon})\delta_{-h+\varepsilon} + p_{\varepsilon}\delta_{h+\varepsilon},$$

where  $p_{\varepsilon} \in (0,1)$  is given by

$$p_{\varepsilon} = \frac{e^{-m(h+\varepsilon)}}{e^{-m(-h+\varepsilon)} + e^{-m(h+\varepsilon)}}.$$

On the one hand, we have  $W_2(F_1, F_2) = \varepsilon$  because  $F_2$  is simply deduced from  $F_1$  by a translation of size  $\varepsilon$ . On the other hand, assuming m'(h) > 0 for simplicity (a similar argument holds if m'(h) < 0) and taking  $\varepsilon > 0$  sufficiently small we obtain  $p_{\varepsilon} < \frac{1}{2}$  and

$$\mathcal{W}_2(\mathcal{M}[F_1], \mathcal{M}[F_2]) = \sqrt{\varepsilon^2 + 4h(h-\varepsilon)\left(\frac{1}{2} - p_\varepsilon\right)} \sim \varepsilon + h \, m'(h)^{1/2} \varepsilon^{1/2},$$

as  $\varepsilon \to 0$  by symmetry and the mean value theorem. In a sense, the leading order term corresponds to the cost of moving a piece of mass  $\frac{1}{2} - p_{\varepsilon}$  from  $-h + \varepsilon$  to h in order to equilibrate the Dirac masses in the transport plan. Since  $m'(h) \neq 0$  this leads to the Lipschitz-discontinuity of the operator  $\mathcal{M}$ .

In [18], this case is ruled out by assuming that the densities are uniformly bounded below on compact intervals. In our case, we will avoid relying on the above quantification in Lemma 2.1 for the contractivity of  $\mathcal{B}$  and will propose a different strategy that we discuss in the sequel.

- 2.3. Brief description of our strategy. As advanced before, our strategy is based on a finer understanding of the iterations of  $\mathcal{T}$  across generations. Specifically, it relies on a suitable reformulation of the solutions  $\{F_n\}_{n\in\mathbb{N}}$  solving the recursion (1.1) for the special quadratic selection function m in (1.5). First, we shall illustrate the case  $\alpha=0$  since our goal is to recapture the special contraction phenomenon behind Lemma 2.1 from our new viewpoint. With that intuition in mind, we will then propose the strategy for  $\alpha>0$ .
- $\diamond$  Case  $\alpha = 0$ . By the previous Corollary 2.2, it is instructive to reformulate the recursion (1.1) in terms of the relative distributions  $\bar{F}_n = F_n/\mathbf{F}_0$ :

$$\boldsymbol{F}_{0}(x)\bar{F}_{n}(x) = \frac{1}{Z} \int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \left| x - \frac{x_{1} + x_{2}}{2} \right|^{2} - \frac{1}{4} |x_{1}|^{2} - \frac{1}{4} |x_{2}|^{2}\right) \frac{\bar{F}_{n-1}(x_{1})\bar{F}_{n-1}(x_{2})}{\int_{\mathbb{R}^{d}} \boldsymbol{F}_{0}(x')\bar{F}_{n-1}(x') dx'} dx_{1} dx_{2},$$

where Z is a n-independent normalizing constant. The quadratic form in the exponential reaches its maximum value at  $(x_1, x_2) = (x/2, x/2)$ . This motivates the change of variables

$$x_1 = \frac{x}{2} + y_1, \quad x_2 = \frac{x}{2} + y_2,$$

which appropriately centers the quadratic form at the minimum and yields the following expression:

$$\mathbf{F}_0(x)\bar{F}_n(x)$$

$$=e^{-\frac{|x|^2}{4}}\frac{1}{Z}\int_{\mathbb{R}^{2d}}\exp\left(-\frac{1}{8}|y_1+y_2|^2-\frac{1}{4}|y_1|^2-\frac{1}{4}|y_2|^2\right)\frac{\bar{F}_{n-1}(\frac{x}{2}+y_1)\bar{F}_{n-1}(\frac{x}{2}+y_2)}{\int_{\mathbb{R}^d}\boldsymbol{F}_0(x')\bar{F}_{n-1}(x')\,dx'}\,dy_1\,dy_2\,. \quad (2.4)$$

Note that by the special scaling in the definition of  $\bar{F}_n$ , the exponential  $e^{-\frac{|x|^2}{4}}$  cancels with the factor  $F_0(x)$  in the left-hand-side (up to a multiplicative constant), so that it remains

$$\bar{F}_n(x) = \int_{\mathbb{R}^{2d}} \widetilde{G}(y_1, y_2) \frac{\bar{F}_{n-1}(\frac{x}{2} + y_1)\bar{F}_{n-1}(\frac{x}{2} + y_2)}{\int_{\mathbb{R}^d} \mathbf{F}_0(x')\bar{F}_{n-1}(x') dx'} dy_1 dy_2, \qquad (2.5)$$

where  $\widetilde{G}$  is a multivariate Gaussian distribution. At this level, we start to observe the contraction phenomenon since the dependence of x on the right hand side of (2.5) has shrunk by a factor 1/2. However, we also note that  $\overline{F}_{n-1}$  appears twice so that it is not evident how to infer the final effect. A

possible strategy to see if there is a quantitative degradation of the dependence on x could be to study the relationship between  $\nabla \log \bar{F}_n$  and  $\nabla \log \bar{F}_{n-1}$ . Specifically, differentiating (2.5) implies:

$$\nabla \log \bar{F}_n(x) = \frac{1}{2} \int_{\mathbb{R}^{2d}} \nabla \log \bar{F}_{n-1} \left( \frac{x}{2} + y_1 \right) \nu_{n-1}(x; dy_1, dy_2) + \frac{1}{2} \int_{\mathbb{R}^{2d}} \nabla \log \bar{F}_{n-1} \left( \frac{x}{2} + y_2 \right) \nu_{n-1}(x; dy_1, dy_2),$$
(2.6)

where the x-dependent probability measures  $\nu_{n-1}(x; dy_1, dy_2)$  on the variables  $(y_1, y_2)$  have the following density with respect to the Lebesgue measure:

$$\frac{\nu_{n-1}(x;dy_1,dy_2)}{dy_1\,dy_2} = \frac{\widetilde{G}(y_1,y_2)\bar{F}_{n-1}(\frac{x}{2}+y_1)\bar{F}_{n-1}(\frac{x}{2}+y_2)}{\int_{\mathbb{R}^{2d}}\widetilde{G}(y_1',y_2')\bar{F}_{n-1}(\frac{x}{2}+y_1')\bar{F}_{n-1}(\frac{x}{2}+y_2')\,dy_1'\,dy_2'}.$$

Therefore, we formally deduce the following non-expansiveness property:

$$\|\nabla \log \bar{F}_n\|_{L^{\infty}(\mathbb{R}^d)} \le \|\nabla \log \bar{F}_{n-1}\|_{L^{\infty}(\mathbb{R}^d)}. \tag{2.7}$$

We emphasize that, in such a rough estimate, we have not used carefully the preservation of the center of mass. This results in the non-expansiveness estimate (2.7), which is conceptually no better than the previous non-expansiveness property (2.3) in the quadratic Wasserstein metric, obtained as in Lemma 2.1 without assuming that the centers of mass are the same. There is one main drawback though: typically  $\nabla \log \bar{F}_n$  are genuinely unbounded unless  $F_n$  and  $F_0$  have the same Gaussian decay for large x (which is much too restrictive). A major point of our work will precisely be to circumvent this unbounded factors. A refinement of the above argument in (2.7) leading to real contractivity would require again tackling more carefully the center of mass. However, we will not address this in the current paper and we refer to Section 8 for some perspectives and future works in this line. For this reason, our method of proof will finally not yield satisfying results in the case  $\alpha = 0$ , but we find this idea illuminating to address the case of a non-trivial effect of selection when  $\alpha > 0$ .

 $\diamond$  Case  $\alpha > 0$ . In this case, our strategy is essentially driven by the previous underlying structure. The first step consists in choosing the appropriate normalization extending the previous definition of  $\bar{F}_n$ . We opted for  $\bar{F}_n = e^m F_n / \mathbf{F}_0$  (which reduces to the above choice when  $\alpha \to 0$ ). There is some freedom here. Actually, the term  $e^m$  is not mandatory, but it is convenient as such for an easier sorting of the various terms. Indeed, we may typically consider the (close-to-optimal) normalization  $F_n/G_{0,\sigma^2}$  for any variance  $\sigma^2$  larger but arbitrarily close to  $\sigma^2_\alpha$  (being  $\sigma^2_\alpha$  the variance (1.13) of the expected equilibrium  $F_\alpha$ ). However, for general selection functions m we do not know the expected equilibrium. To account for a more robust viewpoint, we follow a more robust approach so that we do not "optimize" the normalization.

In a nutshell: we iterate the operator  $\mathcal{T}$  in the recursion for  $F_n$  up to the initial datum and we obtain (modulo a multiplicative constant) a formula of the following type

$$F_n(x) \propto \int_{\mathbb{R}^{(2^{n+1}-2)d}} \exp(-\mathbf{Q}_n(x; \{x_i\})) \prod_i \bar{F}_0(x_i) d\{x_i\}.$$

Here, the index i spans the  $2^{n+1}-2$  members of the genealogical tree with n generations of ancestors from the root (excluded) to the leaves, whilst j only spans the  $2^n$  members at the leaves. In addition,  $\mathbf{Q}_n$  is a quadratic form taking as arguments all traits x and  $\{x_i\}$  (i.e., all the  $2^{n+1}-1$  ancestors in the family chart including the root x). As compared to the case  $\alpha=0$ , we shall see that the minimizers of this high-dimensional quadratic form get shifted towards the origin, as  $\mathbf{Q}_n$  contains additional quadratic contributions  $\frac{\alpha}{2}|x_i|^2$  coming from the selection function (cf. Section 5). As a result of successive backwards changes of variables from leaves to the root very much as in (2.4), the following expression is obtained:

$$F_n(x) \propto \widetilde{\boldsymbol{F}}_n(x) \int_{\mathbb{R}^{(2^{n+1}-2)d}} \exp(-\widetilde{\mathbf{Q}}_n(\{y_i\})) \prod_j \bar{F}_0(\kappa_n x + \Lambda_n^j(\{y_i\})) d\{y_i\}.$$
 (2.8)

Now,  $\widetilde{\mathbf{Q}}_n$  is a new quadratic form and  $\Lambda_n^j$  is a linear map for each leaf j, both taking as arguments all traits  $\{x_i\}$  (root excluded). Then, the dependency upon x is split into two parts: an explicit term  $\widetilde{F}_n(x)$  outside the integral (which is proven to converge towards  $F_\alpha$  by construction) and the contribution  $\kappa_n x$  inside  $\overline{F}_0$  for an appropriate sequence  $\{\kappa_n\}_{n\in\mathbb{N}}\subseteq\mathbb{R}_+$ . The major observation here is that  $\kappa_n\ll 2^{-n}$  because of the strong shift towards the origin under selection.

Formula (2.8) for  $\alpha > 0$  then suggests a stronger form of contraction compared to (2.5) for  $\alpha = 0$ . Indeed, there are still  $2^n$  terms in the product indexed by j, but the contraction parameter  $\kappa_n$  appears to decay fast enough to compensate them. To make this argument quantitative, let us differentiating (2.8) again to obtain

$$\nabla \log F_n(x) = \nabla \log \widetilde{\boldsymbol{F}}_n(x) + \kappa_n \sum_{j} \int_{\mathbb{R}^{(2^{n+1}-2)d}} \nabla \log \bar{F}_0(\kappa_n x + \Lambda_n^j(\{y_i\})) \,\nu_n(x; d\{y_i\}), \qquad (2.9)$$

where the x-dependent probability measures  $\nu_n(x; d\{y_i\})$  on the variables  $\{y_i\}$  have the following density with respect to the Lebesgue measure:

$$\frac{\nu_n(x; d\{y_i\})}{d\{y_i\}} = \frac{\exp(-\widetilde{\mathbf{Q}}_n(\{y_i\})) \prod_j \bar{F}_0(\kappa_n x + \Lambda_n^j(\{y_i\}))}{\int_{\mathbb{R}^{(2^{n+1}-2)d}} \exp(-\widetilde{\mathbf{Q}}_n(\{y_i'\})) \prod_j \bar{F}_0(\kappa_n x + \Lambda_n^j(\{y_i'\})) d\{y_i'\}}$$

If  $\alpha=0$ , then the sequence of parameters degenerates to  $\kappa_n=2^{-n}$  so that (2.9) amounts to the above non-expansiveness property (2.7). However, if  $\alpha>0$ , precisely the fact that  $\kappa_n\ll 2^{-n}$  would infer that the large sum over j consisting of  $2^n$  terms has a vanishing contribution, bounded by  $\kappa_n 2^n \|\nabla \log \bar{F}_0\|_{L^{\infty}(\mathbb{R}^d)}$ . As mention before, there is a drawback: the factors  $\nabla \log \bar{F}_0$  are not really bounded. Then, we shall require more precise estimates for large values of  $\{y_i\}$  in the high-dimensional integral, which are guaranteed by virtue of the strong enough decay of  $\kappa_n$ .

#### 3. Preliminaries

In this part, we collect some necessary preliminary tools and results that will be used later on.

- 3.1. Perfect binary trees. In this paper, we shall systematically use indices i that do not range on discrete intervals  $\{1,\ldots,n\}$  but rather on the vertices of a specific type of trees, which are called perfect binary trees, see [19, 22]. A general definition for trees could be conducted in terms of directed graphs consisting of a finite set of vertices V and a binary relation  $E \subseteq V \times V$  that determines the directed edges connecting a vertex to another vertex. Specifically, a directed graph (V, E) is called a (finite ordered rooted) tree and is denoted T = (V, E) if it does not contain neither loops nor cycles, there is a designated vertex  $\emptyset \in V$  (the root) that is not the target vertex of any edge, and any other vertex is connected with the root through a unique directed path. Since we will restrict to binary trees where each vertex is connected to at most two other vertices, we will simplify our notation and will identify trees with their universal address system, see [19, Section 11.3.3]. Namely, a binary tree T consists in a finite subset of words  $T \subseteq W_2 := \bigcup_{k=0}^{\infty} \{1,2\}^k$  with letters in the alphabet  $\{1,2\}$  that verifies:
  - (1)  $\emptyset \in \mathsf{T}$ .
  - (2) If  $i1 \in \mathsf{T}$  or  $i2 \in \mathsf{T}$  for some word  $i \in \mathsf{W}_2$ , then  $i \in \mathsf{T}$ .

This implies that the root is the empty word, and T is stable under chopping letters on the right of its words. Given a word  $i \in T$ , we denote its length (number of letters) or *height* in the tree by |i|. In particular,  $|\emptyset| = 0$ . A binary tree T is said to be *perfect* if, in addition, the following properties hold:

- (3) Given a word  $i \in W_2$ , then  $i1 \in T$  if, and only if,  $i2 \in T$ .
- (4) There exists  $n \in \mathbb{N}$  such that  $|i| \leq n$  for every  $i \in \mathsf{T}$  and

$$\#\{i \in \mathsf{T} : |i| = n\} = 2^n.$$

This implies that, except for the root, words in T appear in couples  $\{i1,i2\}$  and paths in the tree arising from the root achieve the same maximal height n. The above four conditions determine a unique tree, which we call the *perfect binary tree with height*  $n \in \mathbb{N}$  and we denote by  $T^n$ . See Figure 1 for a graphical representation of the perfect binary tree  $T^3$  of height 3. Binary trees are often used to describe the different generations of offsprings after a given individual. In our case, we shall make a reverse use of trees. Namely, a binary tree will represent the *family tree* or *pedigree chart* of an individual, consisting of the different generations of *ancestors* of such an individual. For this purpose, we shall establish the following terminology:

(1) (**Leaves**) We say that a vertex  $i \in \mathsf{T}^n$  is a leaf of the tree if |i| = n. The set of all leaves of  $\mathsf{T}^n$  will be denoted by

$$\mathsf{L}^n := \{ i \in \mathsf{T}^n : |i| = n \}.$$

(2) (**Levels**) We say that a vertex  $i \in \mathsf{T}^n$  is on the m-th level of the tree if |i| = m for some  $1 \le m \le n$ . The set of all vertices on the m-th level will be denoted

$$L_m^n := \{ i \in T^n : |i| = m \}.$$

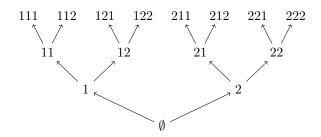


FIGURE 1. Perfect binary tree  $T^3$ 

In particular, note that  $\mathsf{L}^n_0 = \{\emptyset\}$  is the root and  $\mathsf{L}^n_n = \mathsf{L}^n$  are the leaves. For simplicity of notation, we denote the set of vertices of the root-free and leaves-free tree respectively by

$$\mathsf{T}^n_* := \mathsf{T}^n \setminus \{\emptyset\}, \quad \widehat{\mathsf{T}}^n := \mathsf{T}^n \setminus \mathsf{L}^n$$

- (3) (Child) Given a vertex  $i \in \mathsf{T}^n_*$ , then there exists a unique word  $j \in \mathsf{W}_2$  such that i = j1 or i = j2. Such a vertex j is called the *child* of i and will be denoted by  $\mathsf{c}(i) = j$ .
- (4) (Parents) Given a vertex  $\in \widehat{\mathsf{T}}^n$ , we define the *parents* of i as the subset of all vertices that have the same common child i, that is,

$$P(i) := \{i1, i2\}.$$

(5) (Mate) Given a vertex  $i \in \mathsf{T}^n_*$ , we define the *mate* of i and we denote it  $\mathsf{m}(i)$  as the only other vertex in  $\mathsf{T}^n$  that has the same child as i, i.e.,

$$\{i, m(i)\} = \{c(i)1, c(i)2\}.$$

- (6) (Tree order) Given two vertices  $i, j \in \mathsf{T}^n$ , we say that  $i \leq j$  if the associated words are so ordered according to the lexicographical order of the set of words  $\mathsf{W}_2$  with two letters  $\{1, 2\}$ .
- (7) (**Highest common descendant**) Given two vertices  $i, j \in \widehat{\mathsf{T}}^n$ , we define the *highest common descendant* of i and j, and we denote it by  $i \wedge j$ , as

$$i \wedge j := \max\{l \in \mathsf{T}^n : l \le i, l \le j\},$$

where the maximum is considered with respect to above tree (lexicographic) order.

**Remark 3.1** (Tree-indexed variables). Given  $n \in \mathbb{N}$ , we shall identify  $\mathbb{R}^{2(2^n-1)d} \equiv (\mathbb{R}^d)^{\mathsf{T}^n_*}$  and  $\mathbb{R}^{2^n d} \equiv (\mathbb{R}^d)^{\mathsf{L}^n}$ . Specifically, vectors  $\mathbf{y}_n \in \mathbb{R}^{2(2^n-1)d}$  and  $\mathbf{z}_n \in \mathbb{R}^{2^n d}$  will be regarded as indexed families

$$\mathbf{y}_n = (y_i)_{i \in \mathsf{T}_n^n}, \quad \mathbf{z}_n = (z_j)_{j \in \mathsf{L}^n},$$
 (3.1)

where  $y_i \in \mathbb{R}^d$  and  $z_j \in \mathbb{R}^d$  for each  $i \in \mathsf{T}^n_*$  and  $j \in \mathsf{L}^n$ .

3.2. Gaussian solutions. In this part, we compute particular solutions of the time-discrete problem (1.1) and the associated eigenproblem (1.11). As we advanced before, we shall exploit the explicit algebraic structure imposed by G (given by the Gaussian mixing kernel (1.4)) and m (given by the quadratic selection function (1.5)). Namely, explicit Gaussian solution will be obtained. We recall first the following stability property of Gaussians under convolutions.

Lemma 3.2 (Stability of Gaussians). The following relation holds true

$$G_{\mu_1,\sigma_1^2} * G_{\mu_2,\sigma_2^2} = G_{\mu_1+\mu_2,\sigma_1^2+\sigma_2^2},$$
 (3.2)

for any couple of means  $\mu_1, \mu_2 \in \mathbb{R}^d$  and variances  $\sigma_1^2, \sigma_2^2 > 0$ .

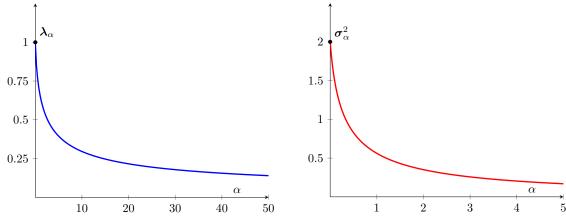
Using the above result, we can obtain the following explicit evaluation of the operator  $\mathcal{T}$  in (1.2) over the class of Gaussian functions.

**Lemma 3.3** (Evaluation on Gaussians). Consider any  $\mu \in \mathbb{R}^d$  and  $\sigma^2 \in \mathbb{R}_+^*$ , then

$$\mathcal{T}[G_{\mu,\sigma^2}] = m_* G_{\mu_*,\sigma^2},$$

where the parameters  $m_*$ ,  $\mu_*$  and  $\sigma_*^2$  are given by:

$$m_* = \frac{e^{-\frac{1}{2} \frac{\alpha |\mu|^2}{1 + \alpha(1 + \frac{\sigma^2}{2})}}}{(1 + \alpha(1 + \frac{\sigma^2}{2}))^{d/2}}, \quad \mu_* = \frac{\mu}{1 + \alpha(1 + \frac{\sigma^2}{2})}, \quad \sigma_*^2 = \frac{1 + \frac{\sigma^2}{2}}{1 + \alpha(1 + \frac{\sigma^2}{2})}.$$
 (3.3)



- (a) The malthusian exponent  $\lambda_{\alpha}$  against parameter  $\alpha$  for d=1.
- (b) The equilibrium variance  $\sigma_{\alpha}^2$  against parameter  $\alpha$ .

FIGURE 2. Plot of eigenvalue  $\lambda_{\alpha}$  and variance at equilibrium  $\sigma_{\alpha}^2$  against  $\alpha$  for d=1. The case  $\alpha=0$  corresponds to absence of selection, which is conservative ( $\lambda_0=1$ ), and the variance at equilibrium  $\sigma_0^2$  is twice the variance of the mixing kernel in  $\mathcal{B}$ .

*Proof.* On the one hand, note that by Lemma 3.2

$$\mathcal{B}[G_{\mu,\sigma^2}](x) = \left(G\left(\frac{\cdot}{2}\right) * G_{\mu,\sigma^2} * G_{\mu,\sigma^2}\right)(2x) = G_{\mu,1+\frac{\sigma^2}{2}}(x),$$

for each  $x \in \mathbb{R}^d$ . Therefore, by definition of  $\mathcal{T}$  in (1.2) we obtain

$$\mathcal{T}[G_{\mu,\sigma^2}](x) = e^{-\frac{\alpha}{2}|x|^2} G_{\mu,1+\frac{\sigma^2}{2}}(x) = \frac{1}{(2\pi(1+\frac{\sigma^2}{2}))^{d/2}} \exp\left(-\frac{\alpha}{2}|x|^2 - \frac{1}{2}\frac{1}{1+\frac{\sigma^2}{2}}|x-\mu|^2\right),$$

for each  $x \in \mathbb{R}^d$ . By completing the square inside the exponential, we conclude our result.

Consequently, the following explicit Gaussian solution of the eigenproblem (1.11) is found.

**Proposition 3.4** (Gaussian solution of the eigenproblem). Assume that  $\alpha \in \mathbb{R}_+^*$ , then the eigenproblem (1.11) has a unique Gaussian solution  $(\lambda_{\alpha}, \mathbf{F}_{\alpha})$ , determined by the relation (1.12) in Theorem 1.1.

*Proof.* We look for  $\lambda \in \mathbb{R}_+^*$ ,  $\mu \in \mathbb{R}^d$  and  $\sigma^2 \in \mathbb{R}_+^*$  such that  $\lambda G_{\mu,\sigma^2} = \mathcal{T}[G_{\mu,\sigma^2}]$ . By Lemma 3.3 and bearing in mind parameters  $m_*$ ,  $\mu_*$  and  $\sigma_*^2$  in (3.3) we obtain that  $(\lambda, \mu, \sigma^2)$  must solve the equations:

$$\lambda = \frac{e^{-\frac{1}{2}\frac{\alpha|\mu|^2}{1+\alpha(1+\frac{\sigma^2}{2})}}}{(1+\alpha(1+\frac{\sigma^2}{2}))^{d/2}}, \quad \mu = \frac{\mu}{1+\alpha(1+\frac{\sigma^2}{2})}, \quad \sigma^2 = \frac{1+\frac{\sigma^2}{2}}{1+\alpha(1+\frac{\sigma^2}{2})}.$$

Hence, the only solution is given by  $\mu = 0$  and  $\lambda$  and  $\sigma^2$  are determined by (1.12) and (1.13).

Remark 3.5 (Dependency on selection). The eigenvalue  $\lambda_{\alpha} \in \mathbb{R}_{+}^{*}$  and the variance  $\sigma_{\alpha}^{2} \in \mathbb{R}_{+}^{*}$  determined by the relations (1.12) and (1.13) for the the unique Gaussian solution  $(\lambda_{\alpha}, \mathbf{F}_{\alpha})$  to the eigenproblem (1.11) are monotonically decreasing with the selection coefficient  $\alpha$ . Namely, the larger  $\alpha$ , the smaller  $\sigma_{\alpha}^{2}$  and  $\lambda_{\alpha}$ . Indeed, we obtain (see also Figure 2)

$$\begin{array}{ccc} \alpha \nearrow \infty & \Longrightarrow & \boldsymbol{\sigma}_{\alpha}^{2} \searrow 0, & \boldsymbol{\lambda}_{\alpha} \searrow 0 \\ \alpha \searrow 0 & \Longrightarrow & \boldsymbol{\sigma}_{\alpha}^{2} \nearrow 2, & \boldsymbol{\lambda}_{\alpha} \nearrow 1. \end{array}$$

Therefore, if selection is strong, then  $\mathbf{F}_{\alpha}$  is very concentrated around the origin and, if selection ceases, then  $\mathbf{F}_{\alpha}$  is twice as spread as the mixing kernel G in (1.4), a famous result in quantitative genetics, see e.g. [8]. Note that  $\lambda_{\alpha} < 1$  for any  $\alpha \in \mathbb{R}_{+}^{*}$  and, consequently, the special steady solutions  $F_{n} = \lambda_{\alpha}^{n} \mathbf{F}_{\alpha}$  coming from ansatz (1.10) always get extinct for large n. This is an artificial consequence of our parameter reduction in Appendix A. In particular, if we maintain parameter  $\beta$  in (A.4) coming from the ratio between birth and mortality rates, then the above special steady solution only extincts if  $\beta < (1 + \alpha(1 + \frac{\sigma_{\alpha}^{2}}{2}))$ .

**Remark 3.6** (Eigenproblem with  $\alpha = 0$ ). The same ideas as in Proposition 3.4 yield the Gaussian solutions to the eigenproblem (1.11) in the absence of selection (i.e.,  $\alpha = 0$ ). However, there is no longer uniqueness due to the translation invariance of  $\mathcal{B}$ . Indeed, we obtain the Gaussian solutions:

$$\lambda = \lambda_0 = 1, \quad F = \mathbf{F}_0(\cdot - \mu) = G_{\mu,2},$$

for any  $\mu \in \mathbb{R}^d$ . Indeed, it is a consequence of the contraction property stated in Section 2.1 that these are all possible generic solutions (not only Gaussian). Specifically, by the conservation of mass (2.1) we obtain that the only possible eigenvalue is  $\lambda = 1$ . Then, the eigenproblem (1.11) reduces to a fixed point equation for  $\mathcal{B}$ . We can then conclude by Corollary 2.2.

We emphasize that the above presents a crucial difference of behavior between the nonlinear problem (1.11) with sexual reproduction and the analogous linear version with asexual reproduction operator (1.7). Namely, whilst in the former case there exists nontrivial solutions in the absence of selection, in the latter case it can be seen that such a nontrivial solution does not exists (e.g., through Fourier arguments). This suggests that whilst in the linear problem both reproduction and selection stay balanced, in the nonlinear problem reproduction appears to dominates selection structurally.

Finally, we note that by iteration of the operator  $\mathcal{T}$  and Lemma 3.3 we can compute the explicit form of solutions of the time-discrete problem (1.1) issued at Gaussian initial data. In fact, the Gaussian structure is preserved along generations, although mass, mean and variance are modified.

**Proposition 3.7** (Gaussian solutions of the time-discrete problem). Consider any  $m_0 \in \mathbb{R}_+^*$ ,  $\mu_0 \in \mathbb{R}^d$  and  $\sigma_0^2 \in \mathbb{R}_+$ , and set the Gaussian initial datum  $F_0 = m_0 G_{\mu_0, \sigma_0^2}$ . Hence, the solution  $\{F_n\}_{n \in \mathbb{N}}$  to the time-discrete problem (1.1) takes the form

$$F_n = m_n G_{\mu_n, \sigma_n^2}, \tag{3.4}$$

for  $n \in \mathbb{N}$ , where the parameters  $m_n \in \mathbb{R}_+^*$ ,  $\mu_n \in \mathbb{R}^d$  and  $\sigma_n^2 \in \mathbb{R}_+^*$  are governed by the recursions:

$$m_n = m_{n-1} \frac{e^{-\frac{1}{2} \frac{\alpha |\mu_{n-1}|^2}{1+\alpha \left(1 + \frac{\sigma_{n-1}^2}{2}\right)}}}{\left(1 + \alpha \left(1 + \frac{\sigma_{n-1}^2}{2}\right)\right)^{d/2}}, \quad \mu_n = \frac{\mu_{n-1}}{1 + \alpha \left(1 + \frac{\sigma_{n-1}^2}{2}\right)}, \quad \frac{1}{\sigma_n^2} = \alpha + \frac{1}{1 + \frac{\sigma_{n-1}^2}{2}}.$$
 (3.5)

**Lemma 3.8** (Relaxation of variance). Assume that  $\alpha \in \mathbb{R}_+^*$ , consider any  $\sigma_0^2 \in \mathbb{R}_+^*$  and define the sequence  $\{\sigma_n^2\}_{n\in\mathbb{N}}$  by recursion according to the third recursion in (3.5), i.e.

$$\frac{1}{\sigma_n^2} = \alpha + \frac{1}{1 + \frac{\sigma_{n-1}^2}{2}},\tag{3.6}$$

for every  $n \in \mathbb{N}$ . Hence, if  $\sigma_0^2 > \sigma_\alpha^2$  then  $\sigma_n \searrow \sigma_\alpha^2$  as  $n \to \infty$  and, if  $\sigma_0^2 < \sigma_\alpha^2$  then  $\sigma_n \nearrow \sigma_\alpha^2$  as  $n \to \infty$ , where  $\sigma_\alpha^2$  is given by (1.13). In addition, we obtain the convergence rates

$$|\sigma_n^2 - \boldsymbol{\sigma}_\alpha^2| \le C_v \, \boldsymbol{r}_\alpha^n, \tag{3.7}$$

for every  $n \in \mathbb{N}$ , where the constant  $C_v \in \mathbb{R}_+$  depends on  $\sigma_0^2$  and  $\alpha$  (we obtain that  $C_v = 0$  if  $\sigma_0^2 = \sigma_\alpha^2$ ), and the ratio  $\mathbf{r}_\alpha \in (0,1)$  is given by

$$r_{\alpha} := \frac{8}{((2\alpha+3) + \sqrt{(2\alpha+1)^2 + 8\alpha})^2}.$$
 (3.8)

*Proof.* • STEP 1: Monotone convergence of  $\{\sigma_n^2\}_{n\in\mathbb{N}}$ . Consider the function  $f: \mathbb{R}_+ \longrightarrow \mathbb{R}$  given by

$$f(x) = \alpha + \frac{1}{1 + \frac{1}{2x}}, \quad x \in \mathbb{R}_+,$$

and, for simplicity of notation, define

$$x_* := (\sigma_\alpha^2)^{-1}, \quad x_n := (\sigma_n^2)^{-1}, \quad n \in \mathbb{N}.$$
 (3.9)

Notice that  $x_*$  is the unique fixed point of f in  $\mathbb{R}_+$  and  $\{x_n\}_{n\in\mathbb{N}}$  determines the fixed-point iteration of the map f issued at  $x_0 = (\sigma_0^2)^{-1}$ , that is,  $x_n = f(x_{n-1}), n \in \mathbb{N}$ . Our goal is to show that  $\{x_n\}_{n\in\mathbb{N}}$  converges towards  $x_*$  and find convergence rates. Since f is not contractive, we cannot apply the usual

Banach contraction principle and a different argument is provided. First, note that  $x \in \mathbb{R}_+ \mapsto f(x)$  is strictly increasing and  $x \in \mathbb{R}_+^* \mapsto \frac{f(x)}{x}$  is strictly decreasing. Then, we obtain that

$$x < f(x) < x_*,$$
 if  $0 \le x < x_*,$   
 $f(x) = x_*,$  if  $x = x_*,$   
 $x_* < f(x) < x,$  if  $x > x_*,$ 

which implies the aforementioned monotonicity properties of  $\{x_n\}_{n\in\mathbb{N}}$  and, in addition,

$$\min\{x_*, x_0\} \le x_n \le \max\{x_*, x_0\},\tag{3.10}$$

for any  $n \in \mathbb{N}$ . In particular, this yields the monotonic convergence of  $\{x_n\}_{n\in\mathbb{N}}$  towards  $x_*$ , or equivalently, the claimed monotonic convergence of  $\{\sigma_n^2\}_{n\in\mathbb{N}}$  towards  $\sigma_\alpha^2$  as  $n \to \infty$ .

#### • Step 2: Convergence rates.

Second, to compute the convergence rates we shall use the special algebraic structure of function f, which is the restriction to  $\mathbb{R}_+$  of the Möbius function  $M: \mathbb{R} \setminus \{-\frac{1}{2}\} \longrightarrow \mathbb{R}$  given by

$$M(x) := \frac{2(\alpha+1)x + \alpha}{2x+1}, \quad x \in \mathbb{R} \setminus \left\{-\frac{1}{2}\right\}.$$

Note that M has two fixed points:

$$x_{\pm} := \frac{2\alpha + 1 \pm \sqrt{(2\alpha + 1)^2 + 8\alpha}}{4}$$

where we note that  $x_+ = x_* \in \mathbb{R}_+^*$  and  $x_- \in \mathbb{R}^-$ . Then, by definition of  $\{x_n\}_{n \in \mathbb{N}}$  in (3.9) we obtain

$$x_n - x_+ = \frac{2(\alpha + 1)x_{n-1} + \alpha}{2x_{n-1} + 1} - \frac{2(\alpha + 1)x_+ + \alpha}{2x_+ + 1} = \frac{2}{(2x_{n-1} + 1)(2x_+ + 1)}(x_{n-1} - x_+). \tag{3.11}$$

Unfortunately, since  $x_0$  can be chosen arbitrary close to 0, then the lower bound on  $x_{n-1}$  is arbitrary close to zero. Hence, by inspection, one notice that contraction on (3.11) would only be guaranteed for large enough  $\alpha$  depending on  $x_0$ . However, for general  $\alpha$  a similar argument yields

$$x_n - x_- = \frac{2(\alpha + 1)x_{n-1} + \alpha}{2x_{n-1} + 1} - \frac{2(\alpha + 1)x_- + \alpha}{2x_- + 1} = \frac{2}{(2x_{n-1} + 1)(2x_- + 1)}(x_{n-1} - x_-).$$
 (3.12)

By dividing (3.11) by (3.12) and iterating the identity we obtain

$$\frac{x_n - x_+}{x_n - x_-} = \left(\frac{2x_- + 1}{2x_+ + 1}\right)^n \frac{x_0 - x_+}{x_0 - x_-},\tag{3.13}$$

for every  $n \in \mathbb{N}$ . In fact the basis can be restated as follows

$$\frac{2x_-+1}{2x_++1} = \frac{(2\alpha+3) - \sqrt{(2\alpha+1)^2 + 8\alpha}}{(2\alpha+3) + \sqrt{(2\alpha+1)^2 + 8\alpha}} = \frac{8}{((2\alpha+3) + \sqrt{(2\alpha+1)^2 + 8\alpha})^2} \equiv \boldsymbol{r}_{\alpha} < 1.$$

Therefore, using the uniform control (3.10) and (3.13), we conclude that

$$|x_n - x_*| = \frac{|x_n - x_-|}{|x_0 - x_-|} |x_0 - x_*| \, \boldsymbol{r}_{\alpha}^n \le \frac{\max\{x_*, x_0\} - x_-}{\min\{x_*, x_0\} - x_-} |x_0 - x_*| \, \boldsymbol{r}_{\alpha}^n, \tag{3.14}$$

for every  $n \in \mathbb{N}$ . Finally, notice that

$$|\sigma_n^2 - \sigma_\alpha^2| = \sigma_n^2 \sigma_\alpha^2 |x_n - x_*| \le \max\{\sigma_\alpha^2, \sigma_0^2\}^2 |x_n - x_*|,$$
 (3.15)

for every  $n \in \mathbb{N}$ . Hence, joining (3.14), (3.15) along with (3.10) ends the proof.

In particular, for Gaussian initial data we recover an explicit particular case of the general relaxation result in Theorem 1.1.

Corollary 3.9 (Relaxation of Gaussian solutions). Assume that  $\alpha \in \mathbb{R}_+^*$ , consider any  $m_0 \in \mathbb{R}_+$ ,  $\mu_0 \in \mathbb{R}^d$  and  $\sigma_0^2 \in \mathbb{R}_+$ , and set the Gaussian initial datum  $F_0 = m_0 G_{\mu_0,\sigma_0}$ . Hence, the Gaussian solution  $\{F_n\}_{n \in \mathbb{N}}$  to the time-discrete problem (1.1) verifies that the growth rates  $||F_n||_{L^1(\mathbb{R}^d)}/||F_{n-1}||_{L^1(\mathbb{R}^d)}$  relax towards

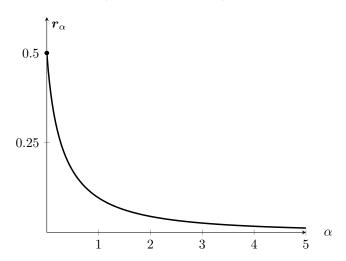


FIGURE 3. The ratio  $r_{\alpha}$  against parameter  $\alpha$ .

 $\lambda_{\alpha}$  and the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  relax towards  $F_{\alpha}$ , where  $(\lambda_{\alpha}, F_{\alpha})$  is the unique Gaussian solution (1.12) to the eigenproblem (1.11) as proved in Proposition 3.4. Specifically, for any  $\varepsilon \in \mathbb{R}_+^*$ 

$$\left| \frac{\|F_n\|_{L^1(\mathbb{R}^d)}}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} - \boldsymbol{\lambda}_{\alpha} \right| \leq C_{\mu,\varepsilon} (\boldsymbol{\lambda}_{\alpha}^{4/d} + \varepsilon)^n + C_v \, \boldsymbol{r}_{\alpha}^n,$$

$$\mathcal{D}_{KL} \left( \frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} \right\| \boldsymbol{F}_{\alpha} \right) \leq C_{\mu,\varepsilon} (\boldsymbol{\lambda}_{\alpha}^{4/d} + \varepsilon)^n + C_v \, \boldsymbol{r}_{\alpha}^n,$$

where  $\mathcal{D}_{KL}$  is the Kullback-Leibler divergence (or relative entropy),  $\mathbf{r}_{\alpha}$  is given by (3.8) in Lemma 3.8,  $C_{\mu,\varepsilon} \in \mathbb{R}_+$  depends on  $\alpha$ ,  $\mu_0$ ,  $\sigma_0^2$  and  $\varepsilon$ , and  $C_v \in \mathbb{R}_+$  depends only  $\alpha$  and  $\sigma_0^2$ . In fact,  $C_{\mu,\varepsilon} = 0$  if  $\mu_0 = 0$  and  $C_v = 0$  if  $\sigma_0^2 = \boldsymbol{\sigma}_{\alpha}^2$ .

*Proof.* First, note that when  $\alpha \in \mathbb{R}_+^*$ , then the parameters  $m_n$ ,  $\mu_n$  and  $\sigma_n^2$  in (3.5) verify

$$\lim_{n\to\infty}\frac{m_n}{m_{n-1}}=\boldsymbol{\lambda}_{\alpha},\quad \lim_{n\to\infty}\mu_n=0,\quad \lim_{n\to\infty}\sigma_n^2=\boldsymbol{\sigma}_{\alpha}^2.$$

In the sequel, we quantify the rates of convergence. Since  $\sigma_n^2$  has already been studied in Lemma 3.8 we shall focus only on  $\mu_n$  and  $m_n$ . Let us fix any arbitrarily small  $\varepsilon \in \mathbb{R}_+^*$ . On the one hand note that

$$\frac{|\mu_n|}{|\mu_{n-1}|} - \lambda_{\alpha}^{2/d} = \frac{1}{1 + \alpha \left(1 + \frac{\sigma_{n-1}^2}{2}\right)} - \frac{1}{1 + \alpha \left(1 + \frac{\sigma_{\alpha}^2}{2}\right)},$$

for every  $n \in \mathbb{N}$ . By the mean value theorem and Lemma 3.8 we obtain

$$\left| rac{|\mu_n|}{|\mu_{n-1}|} - oldsymbol{\lambda}_lpha^{2/d} 
ight| \leq C_v \, oldsymbol{r}_lpha^n,$$

for an appropriate  $C_v \in \mathbb{R}_+^*$  depending on  $\alpha$  and  $\sigma_0^2$ . Since  $\lambda_{\alpha} \in (0,1)$  then d'Alembert's ratio test readily shows that  $\mu_n$  relaxes to zero as a geometric sequence. Indeed, note that

$$|\mu_n| = \left(\frac{|\mu_n|}{|\mu_{n-1}|} - \boldsymbol{\lambda}_{\alpha}^{2/d}\right) |\mu_{n-1}| + \boldsymbol{\lambda}_{\alpha}^{2/d} |\mu_{n-1}| \leq (\boldsymbol{\lambda}_{\alpha}^{2/d} + C_v \, \boldsymbol{r}_{\alpha}^n) |\mu_{n-1}|,$$

for any  $n \in \mathbb{N}$ . By an inductive argument, this yields

$$|\mu_n| \le \prod_{k=1}^n (\boldsymbol{\lambda}_{\alpha}^{2/d} + C_v \, \boldsymbol{r}_{\alpha}^n) |\mu_0| \lesssim C_{\varepsilon} (\boldsymbol{\lambda}_{\alpha}^{2/d} + \varepsilon)^n, \tag{3.16}$$

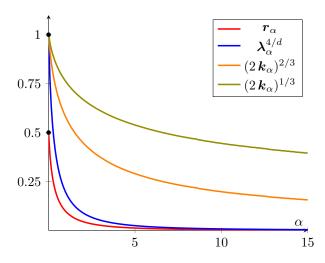


FIGURE 4. Comparison of convergence rates in Theorem 1.1 and Corollary 3.9

for sufficiently large  $C_{\mu,\varepsilon} \in \mathbb{R}^d$ , where we have used that  $\mathbf{r}_{\alpha} \in (0,1)$  to absorb  $C_v \mathbf{r}_{\alpha}^n$  in an  $\varepsilon$ -small term. On the other hand, note that

$$\begin{split} \frac{m_n}{m_{n-1}} - \lambda_{\alpha} &= \frac{e^{-\frac{1}{2}\frac{\alpha|\mu_{n-1}|^2}{1+\alpha\left(1+\frac{\sigma_{n-1}^2}{2}\right)}}}{(1+\alpha(1+\frac{\sigma_{n-1}^2}{2}))^{d/2}} - \frac{1}{(1+\alpha(1+\frac{\sigma_{\alpha}^2}{2}))^{d/2}} \\ &= \frac{e^{-\frac{1}{2}\frac{\alpha|\mu_{n-1}|^2}{2}}}{(1+\alpha(1+\frac{\sigma_{n-1}^2}{2})-1} + \left(\frac{1}{(1+\alpha(1+\frac{\sigma_{n-1}^2}{2}))^{d/2}} - \frac{1}{(1+\alpha(1+\frac{\sigma_{\alpha}^2}{2}))^{d/2}}\right). \end{split}$$

By the mean value theorem,

$$\left| \frac{m_n}{m_{n-1}} - \lambda_{\alpha} \right| \leq \frac{\alpha |\mu_{n-1}|^2}{2(1+\alpha)^{\frac{d}{2}+1}} + \frac{d\alpha}{4} \frac{|\sigma_{n-1}^2 - \sigma_{\alpha}^2|}{(1+\alpha)^{\frac{d}{2}+1}}.$$

Therefore, using contraction of mean (3.16) and contraction of variance (3.7) in Lemma 3.8 entail the rate of convergence for the  $m_n/m_{n-1}$ . Finally, given that both  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  and  $\mathbf{F}_{\alpha}$  are Gaussian functions, the relative entropy can be computed explicitly. Specifically,

$$\mathcal{D}_{\mathrm{KL}}\left(\frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} \| \mathbf{F}_{\alpha}\right) = \frac{|\mu_n|^2}{2 \, \boldsymbol{\sigma}_{\alpha}^2} + \frac{d}{2} \left(\frac{\sigma_n^2}{\boldsymbol{\sigma}_{\alpha}^2} - 1\right) - \frac{d}{2} \log \left(\frac{\sigma_n^2}{\boldsymbol{\sigma}_{\alpha}^2}\right).$$

Hence, using again (3.16) and (3.7) concludes our result.

The proof of Theorem 1.1 for generic non-Gaussian initial data  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  will be the core of this paper and we postpone it to Section 6.

Remark 3.10 (Rough convergence rates). Note that the convergence rates claimed in our main Theorem 1.1 are not sharp compared to those in the previous Corollary 3.9 for the special Gaussian solutions. Specifically, in Corollary 3.9 the sharpest rate is determined by  $\lambda_{\alpha}^{4/d}$  (blue line in Figure 4). However, in Theorem 1.1 the claimed rate is given by  $(2 \mathbf{k}_{\alpha})^{1/3}$  (green line in Figure 4). We believe that this is a technical consequence of the method of proof in our fundamental Lemma 6.3 in Section 6. In this paper, we will not address this "deficiency". Although it might be removed through an appropriate bootstrap argument, unfortunately that does not appear feasible in our case due to the complicatedness of the method.

#### 4. Some properties of the operator $\mathcal{T}$

First, note that for any  $F \in \mathcal{M}_+(\mathbb{R}^d)$  we obtain that  $\mathcal{T}[F] \in W^{k,1}(\mathbb{R}^d) \cap W^{k,\infty}(\mathbb{R}^d)$  for each  $k \in \mathbb{N}$  thanks to the fact that F has finite mass and the Gaussian mixing kernel G in (1.4) is smooth and has bounded and integrable derivatives of any order. In particular, solutions  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1) become instantaneously smooth after the first generation for any generic initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$ . In the sequel, we will derive a quantitative control on the emergence of Gaussian tails

for  $\mathcal{T}[F]$ . In addition, we will quantify the propagation of quadratic and exponential moments for the normalized profile  $\mathcal{T}[F]/\|\mathcal{T}[F]\|_{L^1(\mathbb{R}^d)}$ . Both a priori estimates will be required later in Section 6.

#### 4.1. Emergence of Gaussian tails.

**Lemma 4.1** (Emergence of tails). Assume that  $\alpha \in \mathbb{R}_+^*$ , consider  $F \in \mathcal{M}_+(\mathbb{R}^d)$  and set  $\overline{\sigma}^2, \underline{\sigma}^2 \in \mathbb{R}_+^*$  by

$$\overline{\sigma}^2 := \frac{1}{\alpha}, \quad \underline{\sigma}^2 \in \left(0, \frac{1}{1+\alpha}\right).$$

Then, the following properties are fulfilled:

(1) (Upper control of tails) There exists  $C = C(\alpha, F) > 0$  such that

$$\mathcal{T}[F](x) \le C G_{0,\overline{\sigma}^2}(x),\tag{4.1}$$

for every  $x \in \mathbb{R}^d$ .

(2) (Lower control of tails) There exists  $c = c(\alpha, \underline{\sigma}^2, F) > 0$  such that

$$\mathcal{T}[F](x) \ge c G_{0,\sigma^2}(x),\tag{4.2}$$

for every  $x \in \mathbb{R}^d$ .

(3) (Control of log-derivative) Assume that F is absolutely continuous with respect to the Lebesgue measure and it has Gaussian tail, i.e., there exists  $\sigma^2 \in \mathbb{R}_+^*$  and  $C \in \mathbb{R}_+^*$  such that  $F(x) \leq CG_{0,\sigma^2}$  for every  $x \in \mathbb{R}^d$ . Then, there exists  $D = D(\alpha, \underline{\sigma}^2, \sigma^2, C, F) > 0$  such that

$$|\nabla \log \mathcal{T}[F](x)| \le D(1+|x|),\tag{4.3}$$

for every  $x \in \mathbb{R}^d$ .

*Proof.* First, notice that  $F \in \mathcal{M}_+(\mathbb{R}^d)$  is any generic distribution but  $G \in L^{\infty}(\mathbb{R}^d)$ . Then,  $\mathcal{B}[F] \in L^{\infty}(\mathbb{R}^d)$  and, by definition of  $\mathcal{T}$  in (1.2), we obtain

$$\mathcal{T}[F](x) \le \|\mathcal{B}[F]\|_{L^{\infty}(\mathbb{R}^d)} e^{-\frac{\alpha}{2}|x|^2} = \|\mathcal{B}[F]\|_{L^{\infty}(\mathbb{R}^d)} \left(2\pi\alpha^{-1}\right)^{d/2} G_{0,\overline{\sigma}^2}(x),$$

for every  $x \in \mathbb{R}^d$ . Hence, (4.1) holds for appropriate C > 0. Second, note that

$$\begin{split} \mathcal{T}[F](x) &= \frac{e^{-\frac{\alpha}{2}|x|^2}}{(2\pi)^{d/2} \|F\|_{\mathcal{M}_{+}(\mathbb{R}^d)}} \int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \left|x - \frac{x_1 + x_2}{2}\right|^2\right) F(dx_1) F(dx_2) \\ &= \frac{e^{-\frac{\alpha}{2}|x|^2}}{(2\pi)^{d/2} \|F\|_{\mathcal{M}_{+}(\mathbb{R}^d)}} \int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2}|x|^2 - \frac{1}{8}|x_1 + x_2|^2 + \frac{1}{2}x \cdot (x_1 + x_2)\right) F(dx_1) F(dx_2) \\ &\geq \frac{e^{-\frac{\alpha}{2}|x|^2}}{(2\pi)^{d/2} \|F\|_{\mathcal{M}_{+}(\mathbb{R}^d)}} \int_{\mathbb{R}^{2d}} \exp\left(-\frac{2 + \varepsilon}{4} |x|^2 - \frac{2 + \varepsilon}{8\varepsilon} |x_1 + x_2|^2\right) F(dx_1) F(dx_2) \\ &\geq \frac{e^{-\frac{\alpha}{2}|x|^2}}{(2\pi)^{d/2} \|F\|_{\mathcal{M}_{+}(\mathbb{R}^d)}} \int_{\mathbb{R}^{2d}} \exp\left(-\frac{2 + \varepsilon}{4} |x|^2 - \frac{2 + \varepsilon}{4\varepsilon} |x_1|^2 - \frac{2 + \varepsilon}{4\varepsilon} |x_2|^2\right) F(dx_1) F(dx_2), \end{split}$$

for every  $\varepsilon > 0$ , where in the third line we have used Cauchy–Schwarz's inequality and in the last line we have used Young's inequality. Consequently, we obtain that

$$\mathcal{T}[F](x) \geq (2\pi)^{d/2} \left(\frac{2\varepsilon}{2+\varepsilon}\right)^d \frac{\|G_{0,\frac{2\varepsilon}{2+\varepsilon}}F\|_{\mathcal{M}_+(\mathbb{R}^d)}^2}{\|F\|_{\mathcal{M}_+(\mathbb{R}^d)}} e^{-\frac{\alpha}{2}|x|^2} e^{-\frac{2+\varepsilon}{4}|x|^2},$$

for every  $x \in \mathbb{R}^d$ . Taking  $\varepsilon > 0$  small enough and an appropriate constant c > 0 yields (4.2). Finally, taking derivatives on (1.2) and taking norms yields the following estimate

$$|\nabla \log \mathcal{T}[F](x)| \le (1+\alpha)|x| + \frac{1}{\sqrt{2}||F||_{L^1(\mathbb{R}^d)}} \frac{e^{-\frac{\alpha}{2}|x|^2}}{\mathcal{T}[F](x)} \int_{\mathbb{R}^{2d}} |(x_1, x_2)|G\left(x - \frac{x_1 + x_2}{2}\right) F(dx_1) F(dx_2), \tag{4.4}$$

for every  $x \in \mathbb{R}^d$ . Consider any R > 0 and split the integral in the right hand side of (4.4) into the subsets  $\{(x_1, x_2) \in \mathbb{R}^{2d} : |(x_1, x_2)| \leq \sqrt{2}R|x|\}$  and  $\{(x_1, x_2) \in \mathbb{R}^{2d} : |(x_1, x_2)| > \sqrt{2}R|x|\}$ . For the first

subset, we use the definition (1.2) of the operator  $\mathcal{T}$ . For the second subset, we use the uniform bound of G along with the assumptions on F and the preceding lower bound (4.1) of  $\mathcal{T}[F]$ . Then, we obtain

$$\begin{split} |\nabla \log \mathcal{T}[F](x)| &\leq (1+\alpha+R)|x| + \frac{C^2\underline{\sigma}^d e^{-\frac{\alpha}{2}|x|^2} e^{\frac{1}{2\underline{\sigma}^2}|x|^2}}{\sqrt{2}c||F||_{L^1(\mathbb{R}^d)}} \int_{|(x_1,x_2)| > \sqrt{2}R|x|} |(x_1,x_2)| e^{-\frac{1}{2\sigma^2}|(x_1,x_2)|^2} \, dx_1 \, dx_2 \\ &= (1+\alpha+R)|x| + \frac{C^2\underline{\sigma}^d e^{-\frac{\alpha}{2}|x|^2} e^{\frac{1}{2\underline{\sigma}^2}|x|^2}}{\sqrt{2}c||F||_{L^1(\mathbb{R}^d)}} \int_{\sqrt{2}R|x|}^{+\infty} r^{2d} e^{-\frac{1}{2\sigma^2}r^2} \, dr \\ &\leq (1+\alpha+R)|x| + \frac{C^2\underline{\sigma}^d e^{-\frac{\alpha}{2}|x|^2} e^{\frac{1}{2\underline{\sigma}^2}|x|^2}}{\sqrt{2}c||F||_{L^1(\mathbb{R}^d)}} \left(\frac{8d-4}{e}\right)^{d-\frac{1}{2}} \sigma^{2d-1} \int_{\sqrt{2}R|x|}^{+\infty} r e^{-\frac{1}{4\sigma^2}r^2} \, dr \\ &= (1+\alpha+R)|x| + \frac{\sqrt{2}C^2\underline{\sigma}^d \sigma^{2d+1}}{c||F||_{L^1(\mathbb{R}^d)}} \left(\frac{8d-4}{e}\right)^{d-\frac{1}{2}} e^{-\frac{1}{2}\left(\alpha+\frac{R^2}{\sigma^2}-\frac{1}{\underline{\sigma}^2}\right)|x|^2}, \end{split}$$

for every  $x \in \mathbb{R}^d$ , where in the third line we have used the inequality

$$r^{2d-1} \le \left(\frac{2d-1}{e}\right)^{d-\frac{1}{2}} \left(\frac{\sigma}{\varepsilon}\right)^{2d-1} \exp\left(\frac{\varepsilon}{2\sigma^2}r^2\right),$$

for every r > 0 and  $\varepsilon > 0$  in the particular case of  $\varepsilon = 1/2$ . Taking R > 0 large enough ends the proof.

The previous result can be iterated to obtain similar properties for solutions  $\{F_n\}_{n\in\mathbb{N}}$  to the timediscrete problem (1.1) issued at a generic initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$ . In particular, we note that Gaussian tails emerge instantaneously after the first generation n=1 and linear growth of the logderivative is guaranteed after the second generation n=2.

Corollary 4.2 (Propagation of tails). Assume that  $\alpha \in \mathbb{R}_+^*$ , consider the solution  $\{F_n\}_{n\in\mathbb{N}}$  of (1.1) issued at a generic initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  and set  $\overline{\sigma}_1^2, \underline{\sigma}_1^2 \in \mathbb{R}_+^*$  by

$$\overline{\sigma}_1^2 := \frac{1}{\alpha}, \quad \underline{\sigma}_1^2 \in \left(0, \frac{1}{1+\alpha}\right).$$

Define the associated sequences of variances  $\{\overline{\sigma}_n^2\}_{n\in\mathbb{N}}$  and  $\{\underline{\sigma}_n^2\}_{n\in\mathbb{N}}$  by the following recursive relations

$$\frac{1}{\overline{\sigma}_{n+1}^2} := \alpha + \frac{1}{1 + \frac{\overline{\sigma}_n^2}{2}}, \quad \frac{1}{\underline{\sigma}_{n+1}^2} := \alpha + \frac{1}{1 + \frac{\underline{\sigma}_n^2}{2}}, \tag{4.5}$$

for every  $n \in \mathbb{N}$ . Then, the following properties are fulfilled:

(1) (Upper control of tails of  $F_n$ ) There exist  $C_n = C_n(\alpha, F_0) > 0$  such that

$$F_n(x) \le C_n G_{0,\overline{\sigma}^2}(x), \tag{4.6}$$

for each  $x \in \mathbb{R}^d$  and every  $n \in \mathbb{N}$ . (2) (Lower control of tails of  $F_n$ ) There exist  $c_n = c_n(\alpha, \underline{\sigma}_1^2, F_0) > 0$  such that

$$F_n(x) \ge c_n G_{0,\sigma_n^2}(x),\tag{4.7}$$

for each  $x \in \mathbb{R}^d$  and every  $n \in \mathbb{N}$ .

(3) (Control of log-derivative of  $F_n$ ) There exists  $D_n = D_n(\alpha, \underline{\sigma}_1^2, F_0)$  such that

$$|\nabla \log F_n(x)| \le D_n(1+|x|),\tag{4.8}$$

for each  $x \in \mathbb{R}^d$  and every n > 2.

*Proof.* Note that once (4.6) and (4.7) are proved, we can readily apply (4.3) in Lemma 4.1 to  $F = F_{n-1}$ with  $n \geq 2$  (which satisfies the required upper control by a Gaussian function) and we recover the estimate (4.8) for the log-derivative of  $F_n = \mathcal{T}[F_{n-1}]$ . Then, we just focus on proving estimates (4.6) and (4.7) through an inductive argument. First, note that Lemma 4.1 applied to  $F = F_0$  yields (4.6) and (4.7) with n=1. Let us assume that (4.6) and (4.7) hold for some  $n\in\mathbb{N}$  and let us prove it for n+1. Specifically, using the induction hypothesis we obtain that

$$F_{n+1}(x) = e^{-\frac{\alpha}{2}|x|^2} \mathcal{B}[F_n](x)$$

$$\leq \frac{C_n^2}{\|F_n\|_{L^1(\mathbb{R}^d)}} e^{-\frac{\alpha}{2}|x|^2} \left( G\left(\frac{\cdot}{2}\right) * G_{0,\overline{\sigma}_n^2} * G_{0,\overline{\sigma}_n^2} \right) (2x)$$

$$= \frac{C_n^2}{\|F_n\|_{L^1(\mathbb{R}^d)}} e^{-\frac{\alpha}{2}|x|^2} G_{0,1+\frac{\overline{\sigma}_n^2}{2}}(x) \leq C_{n+1} G_{0,\overline{\sigma}_{n+1}^2}(x),$$

for every  $x \in \mathbb{R}^d$  and appropriate  $C_{n+1}$ , where we have used Lemma 3.2 for the stability of Gaussian under convolutions and the definition (4.5) of  $\overline{\sigma}_{n+1}^2$ . This proves (4.6) for n+1 and a similar argument yields the lower estimate (4.7).

By Lemma 3.8 we note that the sequence of variances  $\{\overline{\sigma}_n^2\}_{n\in\mathbb{N}}$  and  $\{\underline{\sigma}_n^2\}_{n\in\mathbb{N}}$  relax towards the asymptotic value  $\sigma_\alpha^2$  with a geometric convergence rate. This is consistent with our main result in Theorem 1.1 and Corollary 1.2. In fact, this suggests that any solution  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1) must relax towards the asymptotic profile  $F_\alpha$ . In addition, recall that for any solution  $(\lambda, F)$  of the eigenproblem (1.11) we recover a particular solution of the time-discrete problem (1.1) via the ansatz (1.10), i.e.,  $F_n = \lambda^n F$ . Hence, we expect that  $(\lambda_\alpha, F_\alpha)$  must indeed be the unique solution to the eigenvalue problem (1.11). However, we are still far from characterizing the full long-term dynamics for the solution  $\{F_n\}_{n\in\mathbb{N}}$  in Theorem 1.1 and the uniqueness result in Corollary 1.2. Namely, the above coefficients  $C_n$  and  $c_n$  contains crucial information about the balance of mass and their values in Corollary 4.2 are not necessarily optimal. Indeed, note that they are given by explicit recursive formulas

$$C_{n+1} = (1 - \alpha \overline{\sigma}_n^2)^{d/2} \frac{C_n^2}{\|F_n\|_{L^1(\mathbb{R}^d)}}, \quad c_{n+1} = (1 - \alpha \overline{\sigma}_n^2)^{d/2} \frac{c_n^2}{\|F_n\|_{L^1(\mathbb{R}^d)}},$$

but they require further knowledge about the behavior of  $||F_n||_{L^1(\mathbb{R}^d)}$ . In the next paragraph, we overcome this important issue of scaling factor by focusing on renormalized profiles, for which we prove uniform propagation of moments.

4.2. Propagation of quadratic and exponential moments. In this section we analyze the propagation of quadratic and exponential moments along the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$ , where  $\{F_n\}_{n\in\mathbb{N}}$  is a solution to (1.1). Before addressing the case of generic initial data  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$ , we illustrate the particular explicit case of Gaussian initial data  $F_0 = m_0 G_{\mu_0, \sigma_0^2}$ .

**Remark 4.3** (Gaussian case). Consider any  $m_0 \in \mathbb{R}_+$ ,  $\mu_0 \in \mathbb{R}^d$  and  $\sigma_0^2 \in \mathbb{R}_+$ , set  $\varepsilon \in \mathbb{R}_+^*$  with  $\varepsilon < \frac{1}{2 \max\{\sigma_0^2, \sigma_\alpha^2\}}$  and consider the solution  $\{F_n\}_{n \in \mathbb{N}}$  to the time-discrete problem (1.1) with initial datum  $F_0 = m_0 G_{\mu_0, \sigma_0^2}$ . Then, by direct calculation we obtain

$$\int_{\mathbb{R}^d} |x|^2 \frac{F_n(x)}{\|F_n\|_{L^1(\mathbb{R}^d)}} dx = |\mu_n|^2 + \sigma_n^2,$$

$$\int_{\mathbb{R}^d} e^{\varepsilon |x|^2} \frac{F_n(x)}{\|F_n\|_{L^1(\mathbb{R}^d)}} dx = \frac{e^{\frac{\varepsilon}{1-2\varepsilon\sigma_n^2}|\mu_n|^2}}{(1-2\varepsilon\sigma_n^2)^{d/2}},$$

Here,  $\{\mu_n\}_{n\in\mathbb{N}}$  and  $\{\sigma_n^2\}_{n\in\mathbb{N}}$  are the mean and variance of  $\{F_n\}_{n\in\mathbb{N}}$  according to formula (3.5) in Proposition 3.7. As studied in the proof of Corollary 3.9 we have the uniform bounds

$$|\mu_n| \le |\mu_0|$$
 and  $\min\{\sigma_0^2, \boldsymbol{\sigma}_\alpha^2\} \le \sigma_n^2 \le \max\{\sigma_0^2, \boldsymbol{\sigma}_\alpha^2\},$ 

for every  $n \in \mathbb{N}$ . Then, we obtain that both the quadratic and exponential moments are uniformly bounded.

In the sequel, we explore the case of generic initial data  $F_0 \in \mathcal{M}_+(\mathbb{R})$ . For simplicity of notation, we define the following conservative operator associated with  $\mathcal{T}$ .

**Definition 4.4** (Normalization of  $\mathcal{T}$ ).

$$\mathcal{S}[F] := \frac{\mathcal{T}[F]}{\|\mathcal{T}[F]\|_{L^1(\mathbb{R}^d)}}, \quad F \in \mathcal{M}_+(\mathbb{R}^d) \setminus \{0\}.$$

Therefore, we obtain the following control of quadratic moments under  $\mathcal{S}$ .

**Lemma 4.5** (Control of quadratic moments). Assume that  $\alpha \in \mathbb{R}_+^*$  and set any parameter  $\eta \in \left(\frac{1}{2(1+\alpha)^2}, 1\right)$ . Then, there exists a constant  $M = M(\alpha, \eta) \in \mathbb{R}_+^*$  such that

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{S}[F](x) \, dx \le M + \eta \int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}},$$

for any measure  $F \in \mathcal{M}_+(\mathbb{R}^d)$ .

Before we proceed with the proof, let us comment on some apparent difficulties. For this discussion, we assume normalized  $F \in \mathcal{P}(\mathbb{R}^d)$  without loss of generality, as we are only concerned with the shape of profiles here. Note that the conservative operator  $\mathcal{S}$  in Definition 4.4 associated with  $\mathcal{T}$  is actually the composition of the conservative operator  $\mathcal{B}$ , and the conservative multiplicative operator  $\mathcal{M}$  in Definition

2.3, namely,  $S = \mathcal{M} \circ \mathcal{B}$ . The former is invariant by translation, whereas the latter is not. However, we have the following inequality for the operator  $\mathcal{M}$ :

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{M}[F](x) \, dx \le \int_{\mathbb{R}^d} |x|^2 F(dx) \,. \tag{4.9}$$

This inequality is intuitively clear: by applying the selection operator  $\mathcal{M}$ , the quadratic moment (that is, the Wasserstein distance to the Dirac distribution at the origin) is non-increasing. In fact, the proof is a straightforward consequence of the following identity:

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{M}[F](x) \, dx - \int_{\mathbb{R}^d} |x|^2 F(dx) = \frac{1}{2} \int_{\mathbb{R}^{2d}} \left( |x|^2 - |y|^2 \right) \left( e^{-m(x)} - e^{-m(y)} \right) \frac{F(dx) F(dy)}{\int_{\mathbb{R}^d} e^{-m(z)} F(dz)} \, .$$

Note that the right-hand-side of this equality is indeed non-positive, provided that m is radially non-decreasing as it is the case for our quadratic choice (1.5). A similar estimate can also be tested on the operator  $\mathcal{B}$ . Specifically, by a simple change of variable the following identity holds true:

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{B}[F](x) \, dx = \operatorname{Var} G + \frac{1}{2} \int_{\mathbb{R}^d} |x|^2 F(dx) + \frac{1}{2} \left| \int_{\mathbb{R}^d} x F(dx) \right|^2. \tag{4.10}$$

Then, it might seem natural to combine these two relationships in order to control the quadratic moments of the composition  $\mathcal{S} = \mathcal{M} \circ \mathcal{B}$ . Suppose, for instance, that the center of mass in the last term of (4.10) can be bounded a priori uniformly on F. Then, the combination of (4.9) and (4.10) yields the contraction property in Lemma 4.5 due to the reduction by half in the size of the quadratic moment of F. Of course, without such a uniform control on the center of mass, the last two terms in (4.10) would contribute together (by Jensen's inequality) with a merely non-expansive factor  $\int_{\mathbb{R}^d} |x|^2 F(dx)$  that is not enough for our purpose (we will iterate Lemma 4.5 on the  $F_n$  later). This a priori estimate of the center of mass should then be derived mainly from the selection component  $\mathcal{M}$ , as the reproduction component  $\mathcal{B}$  is invariant by translation. However, it cannot be derived solely from  $\mathcal{M}$ , because the latter can have a dramatic effect on nearly symmetric distributions. This can be seen on the same configuration as in Example 2.4 above: for a sum of nearly symmetrical Dirac masses, the center of mass is close to zero before selection, and close to one of the endpoints after selection (when h is large). Nevertheless, such a configuration is destroyed by  $\mathcal{B}$ , so that the combination of  $\mathcal{M}$  and  $\mathcal{B}$  is crucial to guarantee the uniform boundedness of the quadratic moments as shown in the argument below.

Proof of Lemma 4.5. Set any  $F \in \mathcal{M}_{+}(\mathbb{R}^{d})$ . First, note that

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{S}[F](x) \, dx = \frac{\int_{\mathbb{R}^{3d}} |x|^2 e^{-m(x)} G\left(x - \frac{x_1 + x_2}{2}\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} dx}{\int_{\mathbb{R}^{3d}} e^{-m(x)} G\left(x - \frac{x_1 + x_2}{2}\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} dx}.$$

Since G is a Gaussian function given by (1.4) and m is a quadratic function given by (1.5), then we can compute explicitly the integrals with respect to x in the numerator and denominator. Specifically, for fixed  $x_1, x_2 \in \mathbb{R}^d$  we define  $\bar{x} := \frac{x_1 + x_2}{2}$  and we obtain

$$\int_{\mathbb{R}^d} |x|^2 e^{-m(x)} G(x - \bar{x}) \, dx = \frac{1}{(1 + \alpha)^{d/2}} \exp\left(-\frac{1}{2} \frac{\alpha}{1 + \alpha} |\bar{x}|^2\right) \left(\frac{1}{1 + \alpha} + \frac{|\bar{x}|^2}{(1 + \alpha)^2}\right),$$

$$\int_{\mathbb{R}^d} e^{-m(x)} G(x - \bar{x}) \, dx = \frac{1}{(1 + \alpha)^{d/2}} \exp\left(-\frac{1}{2} \frac{\alpha}{1 + \alpha} |\bar{x}|^2\right).$$

Therefore, we can write alternatively

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{S}[F](x) \, dx = \frac{1}{1+\alpha} + \frac{1}{(1+\alpha)^2} \frac{\int_{\mathbb{R}^{2d}} |\bar{x}|^2 \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}}{\int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}}.$$
 (4.11)

Our ultimate goal is to bound the second term in (4.11). To this end, we provide a lower bound for the denominator and an upper bound for the numerator in two separate steps.

• STEP 1: Consider an arbitrary  $R_1 \in \mathbb{R}^*_+$  to be determined later. Then, we obtain

$$\int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^{2}\right) \frac{F(dx_{1})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}} \frac{F(dx_{2})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}} 
\geq \int_{|\bar{x}| \leq R_{1}} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^{2}\right) \frac{F(x_{1})}{\|F\|_{L^{1}(\mathbb{R}^{d})}} \frac{F(x_{2})}{\|F\|_{L^{1}(\mathbb{R}^{d})}} dx_{1} dx_{2} 
\geq \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} R_{1}^{2}\right) \left(1 - \int_{|\bar{x}| > R_{1}} \frac{F(dx_{1})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}} \frac{F(dx_{2})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}}\right) 
\geq \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} R_{1}^{2}\right) \left(1 - \frac{1}{2R_{1}^{2}} \int_{\mathbb{R}^{d}} |x|^{2} \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}}\right). \tag{4.12}$$

• Step 2: Consider an arbitrary  $R_2 \in \mathbb{R}_+^*$  to be determined later. Then, we obtain

$$\int_{\mathbb{R}^{2d}} |\bar{x}|^2 \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} = I_1 + I_2,$$

where each term reads

$$I_{1} := \int_{|\bar{x}| \leq R_{2}} |\bar{x}|^{2} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^{2}\right) \frac{F(dx_{1})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}} \frac{F(dx_{2})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}},$$

$$I_{2} := \int_{|\bar{x}| > R_{2}} |\bar{x}|^{2} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^{2}\right) \frac{F(dx_{1})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}} \frac{F(dx_{2})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}}.$$

On the one hand, the first term can be readily estimated by

$$I_1 \le R_2^2 \int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}.$$
 (4.13)

On the other hand, the second term can be estimated by

$$I_{2} \leq \frac{2}{e\gamma} \frac{1+\alpha}{\alpha} \int_{|\bar{x}| > R_{2}} \exp\left(-\frac{1-\gamma}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^{2}\right) \frac{F(dx_{1})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}} \frac{F(dx_{2})}{\|F\|_{\mathcal{M}(\mathbb{R}^{d})}}$$

$$\leq \frac{2}{e\gamma} \frac{1+\alpha}{\alpha} \exp\left(-\frac{1-\gamma}{2} \frac{\alpha}{1+\alpha} R_{2}^{2}\right), \tag{4.14}$$

for any  $\gamma \in (0,1)$ , where we have used the straightforward inequality

$$r^2 \exp\left(-\frac{1}{2}\frac{\alpha}{1+\alpha}r^2\right) \le \frac{2}{e\gamma}\frac{1+\alpha}{\alpha}\exp\left(-\frac{1-\gamma}{2}\frac{\alpha}{1+\alpha}r^2\right),$$

for any  $r \in \mathbb{R}_+$  with  $r = |\bar{x}|$ .

• Step 3: Putting (4.12), (4.13) and (4.14) into (4.11) yields

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{S}[F](x) \, dx \leq \frac{1}{1+\alpha} + \frac{R_2^2}{(1+\alpha)^2} + \frac{2}{e\gamma} \frac{1+\alpha}{\alpha} \frac{\exp\left(\frac{1}{2} \frac{\alpha}{1+\alpha} R_1^2\right) \exp\left(-\frac{1-\gamma}{2} \frac{\alpha}{1+\alpha} R_2^2\right)}{1 - \frac{1}{2R_1^2} \int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}},$$

for any  $R_1, R_2 \in \mathbb{R}_+^*$  and  $\gamma \in (0, 1)$ . We set the values

$$R_1^2 = \frac{1}{2(1-\delta)} \int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}, \qquad R_2^2 = \frac{R_1^2}{1-\gamma},$$

with  $\gamma, \delta \in (0,1)$ . Then, we obtain

$$\int_{\mathbb{R}^d} |x|^2 \mathcal{S}[F](x) \, dx \leq \frac{1}{1+\alpha} + \frac{2}{e\gamma\delta} \frac{1+\alpha}{\alpha} + \frac{1}{2(1-\delta)(1-\gamma)(1+\alpha)^2} \int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}},$$

thus ending the proof by the arbitrariness of  $\gamma, \delta \in (0,1)$ 

By iteration, we then obtain the following contractivity of the variance of generic solutions of (1.1).

Corollary 4.6 (Propagation of quadratic moments). Assume that  $\alpha \in \mathbb{R}_+^*$ , set any parameter  $\eta \in \left(\frac{1}{2(1+\alpha)^2},1\right)$  and consider the solution  $\{F_n\}_{n\in\mathbb{N}}$  of (1.1) issued at a generic initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  with  $\int_{\mathbb{R}^d} |x|^2 F_0(dx) < \infty$ . Then, there exists  $M = M(\alpha,\eta)$  such that

$$\int_{\mathbb{R}^d} |x|^2 \frac{F_n(x)}{\|F_n\|_{L^1(\mathbb{R}^d)}} \, dx \le \frac{M}{1-\eta} + \eta^n \int_{\mathbb{R}^d} |x|^2 \frac{F_0(dx)}{\|F_0\|_{\mathcal{M}(\mathbb{R}^d)}},$$

for any  $n \in \mathbb{N}$ .

Let us note that, on the one hand the factor  $\eta^n$  makes the dependence on the variance of the initial datum  $F_0$  negligible as the amount of generations increases. On the other hand, the constant M does not depend on the chosen initial profile  $F_0$ . This is compatible with our ergodicity property in Theorem 1.1.

**Lemma 4.7** (Control of exponential moments). Assume that  $\alpha \in \mathbb{R}_+^*$ , set  $\varepsilon \in \mathbb{R}_+^*$  so that  $\varepsilon < \frac{\alpha}{2}$  and any parameter  $\chi > \frac{\varepsilon}{1+\alpha} \frac{\varepsilon}{2(1+\alpha-2\varepsilon)(\alpha-2\varepsilon)}$ . Then, there exists a constant  $C = C(\alpha, \varepsilon, \chi)$  such that

$$\int_{\mathbb{R}^d} e^{\varepsilon |x|^2} \mathcal{S}[F](x) \, dx \le C \left\{ 1 + \exp\left(\chi \int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}\right) \right\},\,$$

for any measure  $F \in \mathcal{M}_+(\mathbb{R}^d)$ .

*Proof.* Since G is a Gaussian function given by (1.4) and m is a quadratic function given by (1.5), then we can compute explicitly the integrals with respect to x in the numerator and denominator like in the proof of Lemma 4.5 and we obtain

$$\int_{\mathbb{R}^d} e^{\varepsilon |x|^2} \mathcal{S}[F](x) dx = \left(\frac{1+\alpha}{1+\alpha-2\varepsilon}\right)^{d/2} \frac{\int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \frac{\alpha-2\varepsilon}{1+\alpha-2\varepsilon} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}}{\int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}}, \quad (4.15)$$

where  $\bar{x} = \frac{x_1 + x_2}{2}$ . Again, we provide bounds for the numerator and denominator in (4.15). Regarding the denominator, we recover estimate (4.12) for any  $R_1 \in \mathbb{R}_+^*$ . Thus, we focus on the bound of the numerator

$$\int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \frac{\alpha - 2\varepsilon}{1 + \alpha - 2\varepsilon} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} = I_1 + I_2,$$

where each term reads

$$I_1 := \int_{|\bar{x}| \le R_2} \exp\left(-\frac{1}{2} \frac{\alpha - 2\varepsilon}{1 + \alpha - 2\varepsilon} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}},$$

$$I_2 := \int_{|\bar{x}| > R_2} \exp\left(-\frac{1}{2} \frac{\alpha - 2\varepsilon}{1 + \alpha - 2\varepsilon} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}.$$

On the one hand, note that

$$I_1 \le \exp\left(\frac{\varepsilon R_2^2}{(1+\alpha-2\varepsilon)^2}\right) \int_{\mathbb{R}^{2d}} \exp\left(-\frac{1}{2} \frac{\alpha}{1+\alpha} |\bar{x}|^2\right) \frac{F(dx_1)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \frac{F(dx_2)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}},\tag{4.16}$$

where we have used the mean value theorem applied to the function  $r \in \mathbb{R}_+^* \mapsto \frac{r}{1+r}$  to derive the relation

$$\frac{\alpha}{1+\alpha} - \frac{\alpha-2\varepsilon}{1+\alpha-2\varepsilon} \leq \frac{2\varepsilon}{(1+\alpha-2\varepsilon)^2}.$$

On the other hand.

$$I_2 \le \exp\left(-\frac{1}{2}\frac{\alpha - 2\varepsilon}{1 + \alpha - 2\varepsilon}R_2^2\right). \tag{4.17}$$

Putting (4.12), (4.16) and (4.17) into (4.15) we obtain

and we end the proof by arbitrariness of  $\delta \in (0,1)$ .

$$\int_{\mathbb{R}^d} e^{\varepsilon |x|^2} \mathcal{S}[F](x) \, dx \leq \left(\frac{1+\alpha}{1+\alpha-2\varepsilon}\right)^{d/2} \left\{ \exp\left(\frac{\varepsilon R_2^2}{(1+\alpha-2\varepsilon)^2}\right) + \frac{\exp\left(-\frac{1}{2}\frac{\alpha-2\varepsilon}{1+\alpha-2\varepsilon}R_2^2\right) \exp\left(\frac{1}{2}\frac{\alpha}{1+\alpha}R_1^2\right)}{1-\frac{1}{2R_1^2}\int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}} \right\},$$

for any  $R_1, R_2 \in \mathbb{R}_+^*$ . As in the proof of Lemma 4.5 we choose them as follows

$$R_1^2 = \frac{1}{2(1-\delta)} \int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}}, \qquad R_2^2 = \frac{\alpha}{1+\alpha} \frac{1+\alpha-2\varepsilon}{\alpha-2\varepsilon} R_1^2,$$

where  $\delta \in (0,1)$ . Then, we obtain

$$\begin{split} \int_{\mathbb{R}^d} e^{\varepsilon |x|^2} \mathcal{S}[F](x) \, dx \\ & \leq \left( \frac{1+\alpha}{1+\alpha-2\varepsilon} \right)^{d/2} \left\{ \frac{1}{\delta} + \exp\left( \frac{\alpha}{1+\alpha} \frac{\varepsilon}{(1+\alpha-2\varepsilon)(\alpha-2\varepsilon)} \frac{1}{2(1-\delta)} \int_{\mathbb{R}^d} |x|^2 \frac{F(dx)}{\|F\|_{\mathcal{M}(\mathbb{R}^d)}} \right) \right\}, \end{split}$$

Therefore, small exponential moments can be controlled by quadratic moments. We can then couple Lemma 4.7 and Corollary 4.6 to obtain the propagation of exponential moments in the time-discrete problem (1.1).

Corollary 4.8 (Propagation of exponential moments). Assume that  $\alpha \in \mathbb{R}_+^*$ , set  $\varepsilon \in \mathbb{R}_+^*$  with  $\varepsilon < \frac{\alpha}{2}$  and any parameter  $\eta \in \left(\frac{1}{2(1+\alpha)^2}, 1\right)$ . Consider the solution  $\{F_n\}_{n \in \mathbb{N}}$  of (1.1) issued at a generic initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  with  $\int_{\mathbb{R}^d} |x|^2 F_0(dx) < \infty$ . Then, there exist  $C = C(\alpha, \varepsilon, \eta)$  and  $C' = C'(\alpha, \varepsilon, \eta)$  such that

$$\int_{\mathbb{R}^d} e^{\varepsilon |x|^2} \frac{F_n(x)}{\|F_n\|_{L^1(\mathbb{R}^d)}} dx \le C \left\{ 1 + \exp\left(C' \left(1 + \eta^{n-1} \int_{\mathbb{R}^d} |x|^2 \frac{F_0(dx)}{\|F_0\|_{\mathcal{M}(\mathbb{R}^d)}}\right)\right) \right\},$$

for any  $n \in \mathbb{N}$ .

#### 5. Reformulating the recursion

In this section we find an appropriate reformulation of the recursion for the solutions  $\{F_n\}_{n\in\mathbb{N}}$  of the time-discrete problem (1.1), as already discussed in Section 2.3. Note that by iterating  $\mathcal{T}$  *n*-times we find that  $F_n$  depends on the initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  via a high-dimensional integral parametrized by variables indexed on the binary tree  $\mathsf{T}^n_*$  (representing the traits of the ancestors in the pedigree chart). Specifically, for any initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  we obtain:

#### $\diamond$ Iteration 1:

By the explicit definition of  $\mathcal{T}$  in (1.2) we readily obtain

$$F_1(x) = \frac{1}{\|F_0\|_{\mathcal{M}(\mathbb{R}^d)}} \int_{\mathbb{R}^{2d}} e^{-m(x)} G\left(x - \frac{x_1 + x_2}{2}\right) F_0(x_1) F_0(x_2) d\mathbf{x}_1, \tag{5.1}$$

where we denote  $\mathbf{x}_1 = (x_1, x_2) \in \mathbb{R}^{2d}$ .

#### ♦ Iteration 2

By (5.1) with  $F_0$  replaced by  $F_1$  and using  $F_1 = \mathcal{T}[F_0]$  in the right hand side we get

$$F_{2}(x) = \frac{1}{\|F_{1}\|_{L^{1}(\mathbb{R}^{d})} \|F_{0}\|_{\mathcal{M}(\mathbb{R}^{d})}^{2}} \int_{\mathbb{R}^{6d}} e^{-(m(x)+m(x_{1})+m(x_{2}))} G\left(x - \frac{x_{1} + x_{2}}{2}\right) \times G\left(x_{1} - \frac{x_{11} + x_{12}}{2}\right) G\left(x_{2} - \frac{x_{21} + x_{22}}{2}\right) F_{0}(x_{11}) F_{0}(x_{12}) F_{0}(x_{21}) F_{0}(x_{22}) d\mathbf{x}_{2}, \quad (5.2)$$

where we denote  $\mathbf{x}_2 = (x_1, x_2, x_{11}, x_{12}, x_{21}, x_{22}) \in \mathbb{R}^{6d}$ .

#### $\diamond$ Iteration $n \in \mathbb{N}$ :

By a clear inductive process, we can iterate the operator  $\mathcal{T}$  as many times  $n \in \mathbb{N}$  as needed to recover an explicit dependence of  $F_n$  on the initial datum  $F_0$ . Of course, this generates a high-dimensional integral involving  $2(2^n - 1)$  variables that we can label along the vertices of the perfect binary tree  $T^n$  according to the universal address system notation in Section 3.1. Namely, we obtain

$$F_{n}(x) = \frac{1}{\prod_{m=0}^{n-1} \|F_{m}\|_{L^{1}(\mathbb{R}^{d})}^{2^{n-1-m}}} \times \int_{\mathbb{R}^{2(2^{n}-1)d}} \exp\left(-\sum_{i \in \widehat{\mathsf{T}}^{n}} m(x_{i})\right) \prod_{i \in \widehat{\mathsf{T}}^{n}} G\left(x_{i} - \frac{x_{i1} + x_{i2}}{2}\right) \prod_{i \in \mathsf{L}^{n}} F_{0}(x_{i}) d\mathbf{x}_{n}, \quad (5.3)$$

where we have used the tree-indexed notation  $\mathbf{x}_n = (x_i)_{i \in \mathsf{T}^n_*} \in \mathbb{R}^{2(2^n-1)d}$  in Remark 3.1 and  $x_\emptyset := x$ .

The ultimate goal of this section is to find a suitable change of variables as discussed in Section 2.3. We shall see that the effect of such a change of variable on the above high-dimensional integral in (5.3) will reveal that actual dependence of  $\{F_n\}_{n\in\mathbb{N}}$  on the shape of  $F_0$ . We will see that such a dependence is indeed weak and the initial datum is rapidly "forgotten" across the different levels of the tree  $\mathsf{T}^n$ . This ergodicity property will be crucial in our analysis and will be exploited later in Section 6 to derive our main result in Theorem 1.1.

5.1. A change of variables across the binary tree. The goal of this section is to appropriately reformulate the high-dimensional integral (5.3) by exploiting the special form of the Gaussian mixing kernel G in (1.4) and the quadratic selection function m in (1.5). Namely, we shall derive a suitable change of variables according to a n-step backwards process starting at the leaf variables  $x_j$  with  $j \in L^n$  and ending at the root variable  $x_0 = x$ . More specifically, at each level m in the tree we shall change the reference frame of the variables  $x_i$  indexed with  $i \in L^n_m$  so that the quadratic form involving those variables in the exponential of (5.3) gets appropriately centered at its minimum. By doing so we find that the minimum is located at a contracted value of the variable  $x_{c(i)} \in L^n_{m-1}$  at the level m-1 below, which is indexed by its children c(i) in the binary tree (recall notation in Section 3.1). By appropriately propagating the backwards process across the tree, and by tracking the accumulated contraction of variables, we shall see that the eventual dependence on the root variable x is weakly gradually forgotten as  $n \to \infty$ . This suggests an apparent form of ergodicity, which we will be crucial later in Section 6.

First, for the sake of clarity, we illustrate our method in the simpler case n = 2. Later, we address the general case with n > 2 driven by the same strategy. As proposed in Section 2.3, we define the following choice for the rescaled profiles that will be used along the change of variables. We recall that the term  $e^m$  is not mandatory but it is convenient for an easier sorting of the various terms.

**Definition 5.1** (Rescaled distributions). Consider any  $F \in \mathcal{M}_+(\mathbb{R}^d)$ . Then, we define

$$\bar{F}(dx) := e^{m(x)} \frac{F(dx)}{F_0(x)}, \quad x \in \mathbb{R}^d.$$

By substituting (1.4) and (1.5) into (5.1) and writing the integral in the right hand side in terms of the rescaled function  $\bar{F}_0$  in Definition 5.1, we obtain that

$$(e^{m}F_{2})(x) = \frac{1}{(2\pi)^{3d/2}} \frac{1}{(4\pi)^{2d}} \frac{1}{\|F_{1}\|_{L^{1}(\mathbb{R}^{d})} \|F_{0}\|_{\mathcal{M}(\mathbb{R}^{d})}^{2}}$$

$$\times \int_{\mathbb{R}^{6d}} \exp\left(-\frac{\alpha}{2} \left(|x_{1}|^{2} + |x_{2}|^{2}\right) - \frac{1}{2} \left|x - \frac{x_{1} + x_{2}}{2}\right|^{2}\right)$$

$$\times \exp\left(-\frac{1}{2} \left|x_{1} - \frac{x_{11} + x_{12}}{2}\right|^{2} - \frac{1}{2} \left|x_{2} - \frac{x_{21} + x_{22}}{2}\right|^{2}\right)$$

$$\times \exp\left(-\frac{\alpha_{0}}{2} \left(|x_{11}|^{2} + |x_{12}|^{2} + |x_{21}|^{2} + |x_{22}|^{2}\right)\right) \bar{F}_{0}(x_{11}) \bar{F}_{0}(x_{12}) \bar{F}_{0}(x_{22}) d\mathbf{x}_{2}.$$

$$(5.4)$$

Here, the parameter  $\alpha_0 \in \mathbb{R}_+^*$  has been defined by

$$\alpha_0 := \frac{1}{2} + \alpha,$$

so that quadratic terms in the last line of (5.4) correct the rescaling  $\bar{F}_0$  of  $F_0$ . The new formulation will be obtained by an appropriate change of variables, so that the quadratic forms inside the above exponential get appropriately centered. To do so, notice that this formula involves as many variables  $x_i$  as indices i in the perfect binary tree  $T^2$ . Indeed, we have sorted the different terms in (5.4) in such a way that the second line only involves variables  $x_1$  and  $x_2$  in the first level  $\mathsf{L}^2_1$  of the tree, whilst the third and the fourth line contain the terms involving the variables  $x_{11}$ ,  $x_{12}$ ,  $x_{21}$  and  $x_{22}$  in the second level of the tree  $\mathsf{L}^2_2$  (i.e., the leaves  $\mathsf{L}^2$ ). We will divide the method into two steps. First, we will address the change of the variables indexed by the leaves  $\mathsf{L}^2$ . Second, we will address the corresponding change of the variables indexed by the first level of the tree  $\mathsf{L}^2_1$ .

• STEP 1: Change of variables for  $x_{11}$ ,  $x_{12}$ ,  $x_{21}$  and  $x_{22}$ . Consider some coefficient  $k_1 > 0$  to be determined later and define the change of variables  $x_{11} \to y_{11}$ ,  $x_{12} \to y_{12}$ ,  $x_{21} \to y_{21}$  and  $x_{22} \to y_{22}$  given by

$$x_{11} = k_1 x_1 + y_{11},$$
  $x_{12} = k_1 x_1 + y_{12},$   $x_{21} = k_1 x_2 + y_{21},$   $x_{22} = k_1 x_2 + y_{12}.$ 

On the one hand, using the change of variables in the terms inside the exponential of (5.4) which involve  $x_{11}$  and  $x_{12}$ , we obtain

$$\frac{\alpha_0}{2}|x_{11}|^2 + \frac{\alpha_0}{2}|x_{12}|^2 + \frac{1}{2}\left|x_1 - \frac{x_{11} + x_{12}}{2}\right|^2 \\
= \frac{\alpha_0}{2}|k_1x_1 + y_{11}|^2 + \frac{\alpha_0}{2}|k_1x_1 + y_{12}|^2 + \frac{1}{2}\left|(1 - k_1)x_1 - \frac{y_{11} + y_{12}}{2}\right|^2 \\
= \frac{1}{2}\left(2\alpha_0k_1^2 + (1 - k_1)^2\right)|x_1|^2 + \frac{\alpha_0}{2}|y_{11}|^2 + \frac{\alpha_0}{2}|y_{12}|^2 + \frac{1}{8}|y_{11} + y_{12}|^2 \\
+ \left(\alpha_0k_1 - \frac{1 - k_1}{2}\right)x_1 \cdot y_{11} + \left(\alpha_0k_1 - \frac{1 - k_1}{2}\right)x_1 \cdot y_{12}.$$

On the other hand, using the change of variables in the terms which involve  $x_{21}$  and  $x_{22}$ , we get

$$\begin{split} \frac{\alpha_0}{2}|x_{21}|^2 + \frac{\alpha_0}{2}|x_{22}|^2 + \frac{1}{2}\left|x_2 - \frac{x_{21} + x_{22}}{2}\right|^2 \\ &= \frac{\alpha_0}{2}|k_1x_2 + y_{21}|^2 + \frac{\alpha_0}{2}|k_1x_2 + y_{22}|^2 + \frac{1}{2}\left|(1 - k_1)x_2 - \frac{y_{21} + y_{22}}{2}\right|^2 \\ &= \frac{1}{2}\left(2\alpha_0k_1^2 + (1 - k_1)^2\right)|x_2|^2 + \frac{\alpha_0}{2}|y_{21}|^2 + \frac{\alpha_0}{2}|y_{22}|^2 + \frac{1}{8}|y_{21} + y_{22}|^2 \\ &\quad + \left(\alpha_0k_1 - \frac{1 - k_1}{2}\right)x_2 \cdot y_{21} + \left(\alpha_0k_1 - \frac{1 - k_1}{2}\right)x_2 \cdot y_{22}. \end{split}$$

Notice that one can eliminate the crossed terms in both expressions by choosing

$$k_1 := \frac{1}{1 + 2\alpha_0} = \frac{1}{2(1 + \alpha)}.$$

In that case, adding both terms we obtain that

$$\frac{\alpha_0}{2}|x_{11}|^2 + \frac{\alpha_0}{2}|x_{12}|^2 + \frac{1}{2}\left|x_1 - \frac{x_{11} + x_{12}}{2}\right|^2 + \frac{\alpha_0}{2}|x_{21}|^2 + \frac{\alpha_0}{2}|x_{22}|^2 + \frac{1}{2}\left|x_2 - \frac{x_{21} + x_{22}}{2}\right|^2 \\
= \frac{1}{2}(1 - k_1)|x_1|^2 + \frac{1}{4k_1}|y_{11}|^2 + \frac{1}{4k_1}|y_{12}|^2 - \frac{1}{8}|y_{11} - y_{12}|^2 \\
+ \frac{1}{2}(1 - k_1)|x_2|^2 + \frac{1}{4k_1}|y_{21}|^2 + \frac{1}{4k_1}|y_{22}|^2 - \frac{1}{8}|y_{21} - y_{22}|^2.$$

Putting everything together into (5.4) yields

$$(e^{m}F_{2})(x) = \frac{1}{(2\pi)^{3d/2}} \frac{1}{(4\pi)^{2d}} \frac{1}{\|F_{1}\|_{L^{1}(\mathbb{R}^{d})} \|F_{0}\|_{\mathcal{M}(\mathbb{R}^{d})}^{2}}$$

$$\times \int_{\mathbb{R}^{6d}} \exp\left(-\frac{\alpha_{1}}{2} (|x_{1}|^{2} + |x_{2}|^{2}) - \frac{1}{2} \left|x - \frac{x_{1} + x_{2}}{2}\right|^{2}\right)$$

$$\times \exp\left(-\frac{1}{4k_{1}} (|y_{11}|^{2} + |y_{12}|^{2} + |y_{21}|^{2} + |y_{22}|^{2}) + \frac{1}{8} (|y_{11} - y_{12}|^{2} + |y_{21} - y_{22}|^{2})\right)$$

$$\times \bar{F}_{0}(k_{1}x_{1} + y_{11}) \bar{F}_{0}(k_{1}x_{1} + y_{12}) \bar{F}_{0}(k_{1}x_{2} + y_{21}) \bar{F}_{0}(k_{1}x_{2} + y_{22})$$

$$\times dx_{1} dx_{2} dy_{11} dy_{12} dy_{21} dy_{22},$$

$$(5.5)$$

where the parameter  $\alpha_1 \in \mathbb{R}_+^*$  has been defined by

$$\alpha_1 := 1 - k_1 + \alpha,$$

in order to recombine the initial terms in the second line of (5.4) involving variables indexed by the first level  $L_1^2$  with the new remainders that have appeared from the previous step.

• STEP 2: Change of variables for  $x_1$  and  $x_2$ .

Consider some coefficient  $k_2 > 0$  to be determined later and define the change of variables  $x_1 \to y_1$  and  $x_2 \to y_2$  given by

$$x_1 = k_2 x + y_1, \quad x_2 = k_2 x + y_2.$$

This time, using the change of variables in the terms of the exponential in the second line of (5.5) yields

$$\begin{split} \frac{\alpha_1}{2}|x_1|^2 + \frac{\alpha_1}{2}|x_2|^2 + \frac{1}{2}\left|x - \frac{x_1 + x_2}{2}\right|^2 \\ &= \frac{\alpha_1}{2}|k_2x + y_1|^2 + \frac{\alpha_1}{2}|k_2x + y_2|^2 + \frac{1}{2}\left|(1 - k_2)x - \frac{y_1 + y_2}{2}\right|^2 \\ &= \frac{1}{2}\left(2\alpha_1k_2^2 + (1 - k_2)^2\right)|x|^2 + \frac{\alpha_1}{2}|y_1|^2 + \frac{\alpha_1}{2}|y_2|^2 + \frac{1}{8}|y_1 + y_2|^2 \\ &+ \left(\alpha_1k_2 - \frac{1 - k_2}{2}\right)x \cdot y_1 + \left(\alpha_1k_2 - \frac{1 - k_2}{2}\right)x \cdot y_2. \end{split}$$

Again, we can cancel the crossed term by choosing

$$k_2 := \frac{1}{1 + 2\alpha_1} = \frac{1}{3 + 2\alpha - 2k_1}.$$

Namely, we obtain that

$$\frac{\alpha_1}{2}|x_1|^2 + \frac{\alpha_1}{2}|x_2|^2 + \frac{1}{2}\left|x - \frac{x_1 + x_2}{2}\right|^2$$

$$= \frac{1}{2}(1 - k_2)|x|^2 + \frac{1}{4k_2}|y_1|^2 + \frac{1}{4k_2}|y_2|^2 - \frac{1}{8}|y_1 - y_2|^2.$$

Then, putting everything together into (5.5) yields

$$F_{2}(x) = \frac{1}{(2\pi)^{3d/2}} \frac{1}{(4\pi)^{2d}} \frac{1}{\|F_{1}\|_{L^{1}(\mathbb{R}^{d})} \|F_{0}\|_{\mathcal{M}(\mathbb{R}^{d})}^{2}} e^{-\frac{1+\alpha-k_{2}}{2}|x|^{2}}$$

$$\times \int_{\mathbb{R}^{6d}} \exp\left(-\frac{1}{4k_{2}} (|y_{1}|^{2} + |y_{2}|^{2}) + \frac{1}{8}|y_{1} - y_{2}|^{2}\right)$$

$$\times \exp\left(-\frac{1}{4k_{1}} (|y_{11}|^{2} + |y_{12}|^{2} + |y_{21}|^{2} + |y_{22}|^{2}) + \frac{1}{8} (|y_{11} - y_{12}|^{2} + |y_{21} - y_{22}|^{2})\right)$$

$$\times \bar{F}_{0}(k_{1}k_{2}x + k_{1}y_{1} + y_{11}) \bar{F}_{0}(k_{1}k_{2}x + k_{1}y_{1} + y_{12})$$

$$\times \bar{F}_{0}(k_{1}k_{2}x + k_{1}y_{2} + y_{21}) \bar{F}_{0}(k_{1}k_{2}x + k_{1}y_{2} + y_{22}) d\mathbf{y}_{2},$$

$$(5.6)$$

where we denote again  $\mathbf{y}_2 = (y_1, y_2, y_{11}, y_{12}, y_{21}, y_{22}) \in \mathbb{R}^{6d}$ .

Before stating the main result for general  $n \in \mathbb{N}$ , we collect some natural notation according to the preceding computations, that will be useful here on. First, we define the following sequences of coefficients.

**Definition 5.2** (Coefficients). Consider any  $\alpha \in \mathbb{R}_+$ .

(1) The coefficients  $\{k_n\}_{n\in\mathbb{N}}$  are defined by the recursive formula

$$k_1 := \frac{1}{2(1+\alpha)},$$

$$k_n := \frac{1}{3+2\alpha-2k_{n-1}}, \quad \text{for } n \ge 2.$$
(5.7)

(2) The coefficients  $\{\kappa_n\}_{n\in\mathbb{N}}$  are defined by the recursive formula

$$\kappa_0 := 1, 
\kappa_n := k_1 \cdots k_n, \quad \text{for } n \ge 1.$$
(5.8)

We note that for n=2 the above coefficients  $k_1, k_2$  reduce to those appearing in the previous reformulation (5.6) of the recursion. Since it will be used later, we study the asymptotic behavior of such sequences of coefficients as  $n \to \infty$ .

**Lemma 5.3** (Asymptotic behavior of the coefficients). Consider the sequences  $\{k_n\}_{n\in\mathbb{N}}$  and  $\{\kappa_n\}_{n\in\mathbb{N}}$  in Definition 5.2. Then, the following properties hold true:

(1) (Coefficients  $k_n$ ) The coefficients  $\{k_n\}_{n\in\mathbb{N}}$  are positive numbers and  $k_n \setminus k_\alpha$  as  $n \to \infty$ , where  $k_\alpha \in \mathbb{R}_+^*$  is the smallest root of the equation

$$\frac{1}{3 + 2\alpha - 2 \, \boldsymbol{k}_{\alpha}} = \boldsymbol{k}_{\alpha} \,. \tag{5.9}$$

Specifically,  $\mathbf{k}_{\alpha}$  is given explicitly by formula (1.14) and it is related to the coefficient  $\mathbf{r}_{\alpha}$  in formula (3.8) of Lemma 3.8 by  $\mathbf{r}_{\alpha} = 2 \mathbf{k}_{\alpha}^{2}$ . In addition,

$$k_n - \mathbf{k}_\alpha \le C \, \mathbf{r}_\alpha^{n-1},\tag{5.10}$$

for every  $n \in \mathbb{N}$ , where  $\mathbf{r}_{\alpha}$  is and  $C \in \mathbb{R}_{+}^{*}$  depends only on  $\alpha$ .

(2) (Coefficients  $\kappa_n$ ) The coefficients  $\{\kappa_n\}_{n\in\mathbb{N}}$  are positive numbers and decay geometrically with maximal rate  $\mathbf{k}_{\alpha}$ . Specifically, for every  $\varepsilon \in \mathbb{R}_+^*$  there exists  $C_{\varepsilon} \in \mathbb{R}_+^*$  such that

$$\kappa_n \le C_{\varepsilon} (\mathbf{k}_{\alpha} + \varepsilon)^n, \tag{5.11}$$

for every  $n \in \mathbb{N}$ .

*Proof.* On the one hand, the properties of  $\{\kappa_n\}_{n\in\mathbb{N}}$  readily follow from those of  $\{k_n\}_{n\in\mathbb{N}}$  and the relation (5.8) in Definition 5.2. Specifically, for any given  $q\in\mathbb{N}$  and any  $n\geq q$  we get the decomposition

$$\kappa_n = \left(\prod_{m=0}^q k_m\right) \left(\prod_{m=q+1}^n k_m\right) \le k_1^q k_{q+1}^{n-q} = \frac{k_1^q}{k_{q+1}^q} k_{q+1}^n \le \frac{k_1^q}{k_{q+1}^q} (C \, \boldsymbol{r}_{\alpha}^{q+1} + \boldsymbol{k}_{\alpha})^n.$$

Therefore, taking q sufficiently large so that  $C r_{\alpha}^{q+1} \leq \varepsilon$  we conclude (5.11). We then focus on the properties of  $\{k_n\}_{n\in\mathbb{N}}$  and we use a similar strategy like in Lemma 3.8.

• STEP 1: Monotone convergence of  $\{k_n\}_{n\in\mathbb{N}}$ . Define the following function

$$f(x) := \frac{1}{3 + 2\alpha - 2x}, \quad x \in (-\infty, \frac{3}{2} + \alpha).$$

Then, by (5.7)  $\{k_n\}_{n\in\mathbb{N}}$  obeys the following recursive relation

$$k_n = f(k_{n-1}), \quad n > 1.$$

Let us consider the fixed points  $x_{-} < x_{+}$  of f, i.e,

$$x_{\pm} = \frac{(3+2\alpha) \pm \sqrt{(3+2\alpha)^2 - 8}}{4}.$$

By inspection, it is clear that  $0 < \mathbf{k}_{\alpha} = x_{-} < x_{+} < \frac{3}{2} + \alpha$  and

$$x_{-} < f(x) < x$$
, if  $x \in (x_{-}, x_{+})$ ,

where  $\mathbf{k}_{\alpha}$  is given in (1.14). Since  $k_1 \in (x_-, x_+)$ , then we conclude that  $\{k_n\}_{n \in \mathbb{N}}$  is a well defined, positive and monotonically decreasing sequence contained in the compact interval  $[x_-, x_+]$ . Therefore, it must converge towards some limit  $\ell$ , which is a solution of f(x) = x i.e.,  $\ell \in \{x_-, x_+\}$ . Since the full sequence  $\{k_n\}_{n \in \mathbb{N}}$  is below  $x_+$  and decreasing, we then conclude that  $\ell = x_- = \mathbf{k}_{\alpha}$ . In particular, we obtain

$$\mathbf{k}_{\alpha} < k_n \le k_1, \tag{5.12}$$

for any  $n \in \mathbb{N}$ .

#### • Step 2: Convergence rates.

As for Lemma 3.8, we shall use the special algebraic structure of the function f, which is the restriction to  $\left(-\infty, \frac{3}{2} + \alpha\right)$  of the Möbius function  $M : \mathbb{R} \setminus \left\{\frac{3}{2} + \alpha\right\} \longrightarrow \mathbb{R}$  given by

$$M(x) := \frac{1}{3 + 2\alpha - 2x}, \quad x \in \mathbb{R} \setminus \left\{ \frac{3}{2} + \alpha \right\}.$$

By definition of  $\{k_n\}_{n\in\mathbb{N}}$  in (5.7) we obtain

$$k_n - x_+ = \frac{1}{3 + 2\alpha - 2k_{n-1}} - \frac{1}{3 + 2\alpha - 2x_+} = \frac{2}{(3 + 2\alpha - 2k_{n-1})(3 + 2\alpha - 2x_+)} (k_{n-1} - x_+).$$

Similarly, we obtain

$$k_n - x_- = \frac{1}{3 + 2\alpha - 2k_{n-1}} - \frac{1}{3 + 2\alpha - 2x_-} = \frac{2}{(3 + 2\alpha - 2k_{n-1})(3 + 2\alpha - 2x_-)}(k_{n-1} - x_-).$$

By dividing both expressions and iterating such an identity we get

$$\frac{k_n - x_-}{k_n - x_+} = \left(\frac{3 + 2\alpha - 2x_+}{3 + 2\alpha - 2x_-}\right)^{n-1} \frac{k_1 - x_-}{k_1 - x_+},\tag{5.13}$$

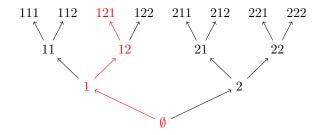


FIGURE 5. Unique path joining the leaf j=121 to the root in the perfect binary tree  $\mathsf{T}^3$ . The corresponding lineage map takes the form  $\Phi_n^{121}(x;\mathbf{y}_3)=\kappa_3x+\kappa_2y_1+\kappa_1y_{12}+y_{121}$ .

for every  $n \in \mathbb{N}$ . In fact, the basis can be restated as follows

$$\frac{3+2\alpha-2x_{+}}{3+2\alpha-2x_{-}} = \frac{3+2\alpha-\sqrt{(3+2\alpha)^{2}-8}}{3+2\alpha+\sqrt{(3+2\alpha)^{2}-8}} = \frac{8}{(3+2\alpha+\sqrt{(3+2\alpha)^{2}-8})^{2}} \equiv \boldsymbol{r}_{\alpha} < 1,$$

where  $r_{\alpha}$  is determined by (3.8) in Lemma 3.8. Therefore, using (5.12) and (5.13) we conclude that

$$k_n - \mathbf{k}_\alpha = \frac{k_1 - x_-}{x_+ - k_1} (x_+ - k_n) \, \mathbf{r}_\alpha^{n-1} \le \frac{k_1 - x_-}{x_+ - k_1} (x_+ - x_-) \, \mathbf{r}_\alpha^{n-1},$$

for any  $n \in \mathbb{N}$ , thus ending the proof.

**Definition 5.4** (Quadratic forms). Consider the sequence  $\{k_n\}_{n\in\mathbb{N}}$  in Definition 5.2. Then, we define the quadratic form  $Q_n = Q_n(\mathbf{y}_n)$  by

$$Q_n(\mathbf{y}_n) := \sum_{m=0}^{n-1} \sum_{i \in \mathsf{L}_m^n} \left( \frac{1}{4k_{n-m}} (|y_{i1}|^2 + |y_{i2}|^2) - \frac{1}{8} |y_{i1} - y_{i2}|^2 \right), \quad \mathbf{y}_n \in \mathbb{R}^{2(2^n - 1)d}, \tag{5.14}$$

where we are using the tree-indexed notation  $\mathbf{y}_n = (y_i)_{i \in \mathsf{T}_*^n} \in \mathbb{R}^{2(2^n-1)d}$  in Remark 3.1.

Again, note that when n=2 the above quadratic form  $Q_2(\mathbf{y}_2)$  reduces to the one inside the exponential of (5.6). We now show the following uniform control of the quadratic forms  $Q_n$  as  $n \to \infty$ .

**Lemma 5.5** (Uniform positive definite quadratic forms). Consider the quadratic form  $Q_n = Q_n(\mathbf{y}_n)$  in Definition 5.4 and define the couple of coefficients  $\beta_{\min}, \beta_{\max} \in \mathbb{R}_+^*$  by

$$eta_{\min} := rac{1+2lpha}{4}, \quad eta_{\max} := rac{1}{4\,m{k}_lpha},$$

where  $\mathbf{k}_{\alpha}$  is given by formula (1.14). Then, we obtain that

$$\beta_{\min} \|\mathbf{y}_n\|^2 \le Q_n(\mathbf{y}_n) \le \beta_{\max} \|\mathbf{y}_n\|^2,$$

where we are using the tree-indexed notation  $\mathbf{y}_n = (y_i)_{i \in \mathsf{T}_+^n} \in \mathbb{R}^{2(2^n-1)d}$  in Remark 3.1.

*Proof.* By virtue of Young's inequality, we obtain the following lower and upper bound for  $Q_n$ 

$$\sum_{m=0}^{n-1} \sum_{i \in \mathsf{L}_m^n} \frac{1}{4} \left( \frac{1}{k_{n-m}} - 1 \right) (|y_{i1}|^2 + |y_{i2}|^2) \le Q_n(\mathbf{y}_n) \le \sum_{m=0}^{n-1} \sum_{i \in \mathsf{L}_m^n} \frac{1}{4} \frac{1}{k_{n-m}} (|y_{i1}|^2 + |y_{i2}|^2),$$

for every  $\mathbf{y}_n = (y_i)_{i \in \mathsf{T}_*^n} \in \mathbb{R}^{2(2^n-1)d}$ . Then, the result follows from Lemma 5.3, which guarantees the uniform control  $\mathbf{k}_{\alpha} < k_{m-n} \le k_1$  for every  $m = 0, \dots, n-1$ .

**Definition 5.6** (Lineage maps). We define the lineage maps  $\Phi_n^j = \Phi_n^j(x; \mathbf{y}_n)$  associated to the leave  $j \in \mathsf{L}^n$  (see Figure 5) as follows

$$\Phi_n^j(x; \mathbf{y}_n) := \kappa_n x + \sum_{m=0}^{n-1} \kappa_m y_{\mathbf{c}^m(j)}, \quad \mathbf{y}_n \in \mathbb{R}^{2(2^n - 1)d},$$
 (5.15)

where  $x \in \mathbb{R}^d$  represents the root value,  $\mathbf{y}_n = (y_i)_{i \in \mathsf{T}_*^n} \in \mathbb{R}^{2(2^n-1)d}$  is represented according to the tree-indexed notation in Remark 3.1, the sequence  $\{\kappa_n\}_{n \in \mathbb{N}}$  is given in Definition 5.2 and  $\mathbf{c} : \mathsf{T}_*^n \to \widehat{\mathsf{T}}^n$  is the map in Section 3.1 that, to any vertex  $i \in \mathsf{T}_*^n$ , it associates its child  $\mathbf{c}(i) \in \widehat{\mathsf{T}}^n$ .

We are now ready to state the main result of this section extending formula (5.6) to any  $n \in \mathbb{N}$ . The starting point is again formula (5.3). In the particular case of Gaussian mixing kernel (1.4) and quadratic selection function (1.5), it takes the explicit form

$$(e^{m}F_{n})(x) = \frac{1}{(2\pi)^{(2^{n}-1)d/2}} \frac{1}{(4\pi)^{2^{n-1}d}} \frac{1}{\prod_{m=0}^{n-1} ||F_{m}||_{L^{1}(\mathbb{R}^{d})}^{2^{n-1-m}}} \times \int_{\mathbb{R}^{2(2^{n}-1)d}} \exp\left(-\sum_{i \in \widehat{\mathsf{T}}^{n}} \left[\frac{\alpha}{2}(|x_{i1}|^{2} + |x_{i2}|^{2}) + \frac{1}{2} \left|x_{i} - \frac{x_{i1} + x_{i2}}{2}\right|^{2}\right]\right) \times \exp\left(-\frac{1}{4} \sum_{j \in \mathsf{L}^{n}} |x_{j}|^{2}\right) \prod_{j \in \mathsf{L}^{n}} \bar{F}_{0}(x_{j}) d\mathbf{x}_{n}, \quad (5.16)$$

for any  $x \in \mathbb{R}^d$  and  $n \in \mathbb{N}$ , where we have used the tree-indexed notation  $\mathbf{x}_n = (x_i)_{i \in \mathbb{T}^n_*} \in \mathbb{R}^{2(2^n-1)d}$  in Remark 3.1, we have set  $x_\emptyset := x$  and we have considered the rescaled distribution  $\bar{F}_0$  according to Definition 5.1. Again, notice that the exponential terms indexed by the leaves  $\mathsf{L}^n$  in the last line of (5.16) have been introduced to appropriately correct the rescaled distribution  $\bar{F}_0$  of  $F_0$ . As a consequence, the factor  $(4\pi)^{-2^{n-1}d}$  arises from the normalization by  $F_0(x_j)$  whilst the factor  $(2\pi)^{-(2^n-1)d/2}$  comes from the repeated products of the Gaussian mixing kernel G. The main result then reads as follows

**Proposition 5.7** (Reformulation of the iterations I). Assume that  $\alpha \in \mathbb{R}_+$  and set any initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$ . Hence, the solution  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1) admits the following form

$$F_n(x) = \frac{e^{-\frac{1+\alpha-k_n}{2}|x|^2}}{(4\pi)^{2^{n-1}d}(2\pi)^{(2^n-1)d/2}} \frac{1}{\prod_{m=0}^{n-1} \|F_m\|_{L^1(\mathbb{R}^d)}^{2^{n-1-m}}} \int_{\mathbb{R}^{2(2^n-1)d}} e^{-Q_n(\mathbf{y}_n)} \prod_{j \in \mathbb{L}^n} \bar{F}_0(\Phi_n^j(x; \mathbf{y}_n)) d\mathbf{y}_n, \quad (5.17)$$

for every  $x \in \mathbb{R}^d$  and  $n \in \mathbb{N}$ , where we are using the tree-indexed notation  $\mathbf{y}_n = (y_i)_{i \in \mathsf{T}^n_*} \in \mathbb{R}^{2(2^n-1)d}$  in Remark 3.1 along with further notation from Definitions 5.1, 5.2, 5.4 and 5.6.

As we anticipated for the particular case n=2 at the beginning of this section, the proof of Proposition 5.7 will be based on the following change of variables

$$x_{i1} = k_{n-m}x_i + y_{i1}, \quad x_{i2} = k_{n-m}x_i + y_{i2},$$

for every  $i \in \mathsf{L}_m^n$ , which we will apply in a backwards way starting at indices  $i \in \mathsf{L}_{n-1}^n$  and ending at  $i = \emptyset$ . Again, the objective of such a change of variables is to appropriately move the reference frame so that the above quadratic forms  $\frac{1}{2} \left| x_i - \frac{x_{i1} + x_{i2}}{2} \right|^2$  inside the exponential of (5.16) get appropriately centered and no crossed terms remain when the selection parts  $\frac{\alpha}{2} (|x_{i1}|^2 + |x_{i2}|^2)$  are taken into account. For simplicity, along the proof we shall restrict to tracking how the various terms inside the exponentials in (5.16) get modified after the change of variables.

Proof of Proposition 5.7.

• STEP 1: Change of variables  $x_{i1} \to y_{i1}$  and  $x_{i2} \to y_{i2}$  for  $i \in L_{n-1}^n$ . We consider the change of variables given by

$$x_{i1} = k_1 x_i + y_{i1}, \quad x_{i2} = k_1 x_i + y_{i2},$$

where  $k_1$  is given by Definition 5.2. The collection of all the terms for  $i \in L_{n-1}^n$  in (5.16) then read

$$\frac{\alpha_0}{2}|x_{i1}|^2 + \frac{\alpha_0}{2}|x_{i2}|^2 + \frac{1}{2}\left|x_i - \frac{x_{i1} + x_{i2}}{2}\right|^2 \\
= \frac{\alpha_0}{2}|k_1x_i + y_{i1}|^2 + \frac{\alpha_0}{2}|k_1x_i + y_{i2}|^2 + \frac{1}{2}\left|(1 - k_1)x_i - \frac{y_{i1} + y_{i2}}{2}\right|^2 \\
= \frac{1}{2}\left(2\alpha_0k_1^2 + (1 - k_1)^2\right)|x_i|^2 + \frac{\alpha_0}{2}|y_{i1}|^2 + \frac{\alpha_0}{2}|y_{i2}|^2 + \frac{1}{8}|y_{i1} + y_{i2}|^2 \\
+ \left(\alpha_0k_1 - \frac{1 - k_1}{2}\right)x_i \cdot y_{i1} + \left(\alpha_0k_1 - \frac{1 - k_1}{2}\right)x_i \cdot y_{i2} \\
= \frac{1}{2}\left(2\alpha_0k_1^2 + (1 - k_1)^2\right)|x_i|^2 + \frac{\alpha_0}{2}|y_{i1}|^2 + \frac{\alpha_0}{2}|y_{i2}|^2 + \frac{1}{8}|y_{i1} + y_{i2}|^2 \\
= \frac{1}{2}(1 - k_1)|x_i|^2 + \frac{1}{4k_1}|y_{i1}|^2 + \frac{1}{4k_1}|y_{i2}|^2 - \frac{1}{8}|y_{i1} - y_{i2}|^2,$$

where we have defined the coefficient  $\alpha_0 \in \mathbb{R}_+^*$  by

$$\alpha_0 := \alpha + \frac{1}{2},$$

to recombine terms in the second and third lines of (5.16), and we have used the following relation between  $k_1$  and  $\alpha_0$  in order to cancel the crossed terms, *i.e.*,

$$k_1 = \frac{1}{1 + 2\alpha_0}.$$

• STEP 2: Change of variables  $x_{i1} \to y_{i1}$  and  $x_{i2} \to y_{i2}$  for  $i \in \mathsf{L}^n_{n-2}$ . Again, we consider the change of variables given by

$$x_{i1} = k_2 x_i + y_{i1}, \quad x_{i2} = k_2 x_i + y_{i2},$$

where  $k_2$  is given by Definition 5.2. Putting together the terms for  $i \in L_{n-2}^n$  in (5.16) and the above  $x_i$ -dependent remainder in the above expression in STEP 1 yields the following term under the above change of variables

$$\begin{split} \frac{\alpha_1}{2}|x_{i1}|^2 + \frac{\alpha_1}{2}|x_{i2}|^2 + \frac{1}{2}\left|x_i - \frac{x_{i1} + x_{i2}}{2}\right|^2 \\ &= \frac{\alpha_1}{2}|k_2x_i + y_{i1}|^2 + \frac{\alpha_1}{2}|k_2x_i + y_{i2}|^2 + \frac{1}{2}\left|(1 - k_2)x_i - \frac{y_{i1} + y_{i2}}{2}\right|^2 \\ &= \frac{1}{2}\left(2\alpha_1k_2^2 + (1 - k_2)^2\right)|x_i|^2 + \frac{\alpha_1}{2}|y_{i1}|^2 + \frac{\alpha_1}{2}|y_{i2}|^2 + \frac{1}{8}|y_{i1} + y_{i2}|^2 \\ &\quad + \left(\alpha_1k_2 - \frac{1 - k_2}{2}\right)x_i \cdot y_{i1} + \left(\alpha_1k_2 - \frac{1 - k_2}{2}\right)x_i \cdot y_{i2} \\ &= \frac{1}{2}\left(2\alpha_1k_2^2 + (1 - k_2)^2\right)|x_i|^2 + \frac{\alpha_1}{2}|y_{i1}|^2 + \frac{\alpha_1}{2}|y_{i2}|^2 + \frac{1}{8}|y_{i1} + y_{i2}|^2 \\ &= \frac{1}{2}(1 - k_2)|x_i|^2 + \frac{1}{4k_2}|y_{i1}|^2 + \frac{1}{4k_2}|y_{i2}|^2 - \frac{1}{8}|y_{i1} - y_{i2}|^2, \end{split}$$

where we have defined again coefficient the  $\alpha_1 \in \mathbb{R}_+^*$  by

$$\alpha_1 := 1 + \alpha - k_1,$$

in order to absorb the above-mentioned  $x_i$ -dependent remainder. In addition, note that we have used again the following relation between  $k_2$  and  $\alpha_1$  to cancel the crossed term, *i.e.*,

$$k_2 = \frac{1}{1 + 2\alpha_1}.$$

• STEP 3: Change of variables  $x_{i1} \to y_{i1}$  and  $x_{i2} \to y_{i2}$  for  $i \in \mathsf{L}^n_{n-3}$ . We consider the change of variables given by

$$x_{i1} = k_3 x_i + y_{i1},$$
  
$$x_{i2} = k_3 x_i + y_{i2},$$

where  $k_3$  is given by Definition 5.2. Putting together the terms for  $i \in L_{n-3}^n$  in (5.16) and the above  $x_i$ -dependent remainder in the above expression in STEP 2 yields the following term under the above change of variables

$$\begin{split} \frac{\alpha_2}{2}|x_{i1}|^2 + \frac{\alpha_2}{2}|x_{i2}|^2 + \frac{1}{2}\left|x_i - \frac{x_{i1} + x_{i2}}{2}\right|^2 \\ &= \frac{\alpha_2}{2}|k_3x_i + y_{i1}|^2 + \frac{\alpha_2}{2}|k_3x_i + y_{i2}|^2 + \frac{1}{2}\left|(1 - k_3)x_i - \frac{y_{i1} + y_{i2}}{2}\right|^2 \\ &= \frac{1}{2}\left(2\alpha_2k_3^2 + (1 - k_3)^2\right)|x_i|^2 + \frac{\alpha_2}{2}|y_{i1}|^2 + \frac{\alpha_2}{2}|y_{i2}|^2 + \frac{1}{8}|y_{i1} + y_{i2}|^2 \\ &\quad + \left(\alpha_2k_3 - \frac{1 - k_3}{2}\right)x_i \cdot y_{i1} + \left(\alpha_2k_3 - \frac{1 - k_3}{2}\right)x_i \cdot y_{i2} \\ &= \frac{1}{2}\left(2\alpha_2k_3^2 + (1 - k_3)^2\right)|x_i|^2 + \frac{\alpha_2}{2}|y_{i1}|^2 + \frac{\alpha_2}{2}|y_{i2}|^2 + \frac{1}{8}|y_{i1} + y_{i2}|^2 \\ &= \frac{1}{2}(1 - k_3)|x_i|^2 + \frac{1}{4k_3}|y_{i1}|^2 + \frac{1}{4k_3}|y_{i2}|^2 - \frac{1}{8}|y_{i1} - y_{i2}|^2, \end{split}$$

where we have defined again coefficient the  $\alpha_2 \in \mathbb{R}_+^*$ 

$$\alpha_2 := 1 + \alpha - k_2$$

in order to absorb the above-mentioned  $x_i$ -dependent remainder. In addition, note that we have used again the following relation between  $k_3$  and  $\alpha_2$  to cancel the crossed term, *i.e.*,

$$k_3 = \frac{1}{1 + 2\alpha_2}.$$

Following a similar recursive process, we readily identify the quadratic form  $Q_n$  in Definition 5.4 in the exponential of the integrand in the final expression (5.17).

• STEP 4: Identifying the lineage maps  $\Phi_n^j$  and the exponential factor. On the one hand, the hierarchy of changes of variables immediately leads to

$$x_j = \Phi_n^j(x; \mathbf{y}_n),$$

for any leaf  $j \in L^n$ , thanks to Definition 5.2 for the coefficients  $\{\kappa_n\}_{n\in\mathbb{N}}$  and Definition 5.6 for the lineage maps  $\Phi_n^j$ . On the other hand, since the Jacobian determinant of each change of variables is 1, no further factors appear during the iterative process except for the exponential x-dependent remainder  $e^{-\frac{1+\alpha-k_n}{2}|x|^2}$  of the last step of the recurrence, which is not absorbed in the quadratic form  $Q_n$ .

5.2. **Probabilistic reinterpretation.** For general purposes, in this part we reformulate the result in Proposition 5.7 in appropriate probabilistic terms. We shall use the following notation.

**Definition 5.8** (High-dimensional normal distributions). Consider any integer  $n \in \mathbb{N}$  and

$$\Sigma_{n}^{m} := \begin{pmatrix} \frac{(2-k_{n-m})k_{n-m}}{1-k_{n-m}}I_{d} & -\frac{k_{n-m}^{2}}{1-k_{n-m}}I_{d} \\ -\frac{k_{n-m}^{2}}{1-k_{n-m}}I_{d} & \frac{(2-k_{n-m})k_{n-m}}{1-k_{n-m}}I_{d} \end{pmatrix},$$
(5.18)

for every m = 0, ..., n-1. Then, we will define the random vector  $\mathbf{Y}_n = (Y_i)_{i \in \mathsf{T}_*^n}$  distributed according to the following multivariate normal distribution

$$G_n(\mathbf{y}_n) := \prod_{m=0}^{n-1} \prod_{i \in \mathbf{L}_n^m} \frac{1}{(2\pi)^d \sqrt{\det(\Sigma_n^m)}} \exp\left(-\frac{1}{2} (y_{i1}^T, y_{i2}^T) (\Sigma_n^m)^{-1} \begin{pmatrix} y_{i1} \\ y_{i2} \end{pmatrix}\right), \quad \mathbf{y}_n \in \mathbb{R}^{2(2^n-1)d}, \tag{5.19}$$

where again we are using the tree-indexed notation  $\mathbf{y}_n = (y_i)_{i \in \mathsf{T}^n_*} \in \mathbb{R}^{2(2^n-1)d}$  in Remark 3.1.

Notice that for every index  $i \in \mathsf{L}_m^n$  with  $m = 0, \ldots, n-1$  the pair  $(Y_{i1}, Y_{i2})$  is normally distributed according to  $\mathcal{N}(0, \Sigma_n^m)$  and  $Y_{i1}$  and  $Y_{i2}$  are certainly correlated. Nevertheless, the vector  $(Y_{i1}, Y_{i2})$  is independent on any other random vector  $Y_j$  for indices  $j \neq i1$  and  $j \neq i2$ . We are now ready to introduce the following probabilistic reformulation of Proposition 5.7.

**Proposition 5.9** (Reformulation of the iterations II). Assume that  $\alpha \in \mathbb{R}_+^*$  and set any initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$ . Hence, the solution  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1) admits the following form

$$F_n(x) = \frac{e^{-\frac{1+\alpha-k_n}{2}|x|^2}}{(2\pi)^{d/2} 2^{2^{n-1}d}} \left( \prod_{m=0}^{n-1} \frac{\left(\frac{4k_{n-m}^2}{1-k_{n-m}}\right)^{2^{m-1}d}}{\|F_m\|_{L^1(\mathbb{R}^d)}^{2^{n-1-m}}} \right) \mathbb{E}\left[ \prod_{j \in \mathsf{L}^n} \bar{F}_0(\Phi_n^j(x; \mathbf{Y}_n)) \right], \tag{5.20}$$

for every  $x \in \mathbb{R}^d$  and  $n \in \mathbb{N}$ , where we are using the high-dimensional random vector  $\mathbf{Y}_n = (Y_i)_{i \in \mathsf{T}^n_*}$  in Definition 5.8 along with further notation from Definitions 5.1, 5.2, 5.4 and 5.6.

*Proof.* By Proposition 5.7 we obtain

$$\begin{split} F_n(x) &= \frac{e^{-\frac{1+\alpha-k_n}{2}|x|^2}}{(4\pi)^{2^{n-1}d}(2\pi)^{(2^n-1)d/2}} \frac{1}{\prod_{m=0}^{n-1} \|F_m\|_{L^1(\mathbb{R}^d)}^{2^{n-1-m}}} \int_{\mathbb{R}^{2(2^n-1)d}} e^{-Q_n(\mathbf{y}_n)} \prod_{j \in \mathbb{L}^n} \bar{F}_0(\Phi_n^j(x;\mathbf{y}_n)) \, d\mathbf{y}_n \\ &= \frac{(2\pi)^{(2^n-1)d/2}}{(4\pi)^{2^{n-1}d}} e^{-\frac{1+\alpha-k_n}{2}|x|^2} \left( \prod_{m=0}^{n-1} \frac{\left(\frac{4k_{n-m}^2}{1-k_{n-m}}\right)^{2^{m-1}d}}{\|F_m\|_{L^1(\mathbb{R}^d)}^{2^{n-1-m}}} \right) \int_{\mathbb{R}^{2(2^n-1)d}} G_n(\mathbf{y}_n) \prod_{j \in \mathbb{L}^n} \bar{F}_0(\Phi_n^j(x;\mathbf{y}_n)) \, d\mathbf{y}_n, \end{split}$$

for every  $x \in \mathbb{R}^d$  and each  $n \in \mathbb{N}$ , where in the last line we have used that the inverse and determinant of the covariance matrices  $\Sigma_n^m$  in (5.18) take the form

$$(\Sigma_n^m)^{-1} := \begin{pmatrix} \frac{2 - k_{n-m}}{4k_{n-m}} I_d & \frac{1}{4} I_d \\ \frac{1}{4} I_d & \frac{2 - k_{n-m}}{4k_{n-m}} I_d \end{pmatrix}, \quad \det(\Sigma_n^m) = \left(\frac{4k_{n-m}^2}{1 - k_{n-m}}\right)^d.$$

Then, the result follows by applying the law of the unconscious statistician (LOTUS) to relate the preceding integral to the expectation of the random variable  $\prod_{j \in \mathbb{L}^n} \bar{F}_0(\Phi_n^j(x; \mathbf{Y}_n))$ .

Remark 5.10 (Non-negligible correlations). Since  $\mathbf{Y}_n$  in Definition 5.8 is normally distributed and the lineage map  $\Phi_n^j(x;\mathbf{y}_n)$  in Definition 5.6 is an affine transformation of  $\mathbf{y}_n$ , then  $(\Phi_n^j(x;\mathbf{Y}_n))_{j\in\mathbb{L}^n}$  are also normally distributed. In addition, since the components of  $\mathbf{Y}_n$  indexed by any couple of indices (i1,i2) are correlated so are all the components of  $(\Phi_n^j(x;\mathbf{Y}_n))_{j\in\mathbb{L}^n}$ . Specifically, a straightforward computation shows that the covariance matrix between any couple  $(i,j)\in\mathbb{L}^n\times\mathbb{L}^n$  of components takes the form

$$\left(\sum_{q=0}^{n-1} \kappa_q^2 \frac{(2-k_{q+1})k_{q+1}}{1-k_{q+1}}\right) I_d, \qquad if \quad |i \wedge j| = n, 
\left(-\kappa_{n-|i \wedge j|-1}^2 \frac{k_{n-|i \wedge j|}^2}{1-k_{n-|i \wedge j|}} + \sum_{q=n-|i \wedge j|}^{n-1} \kappa_q^2 \frac{(2-k_{q+1})k_{q+1}}{1-k_{q+1}}\right) I_d, \quad if \quad 0 < |i \wedge j| < n, 
\left(-\kappa_{n-1}^2 \frac{k_n^2}{1-k_n}\right) I_d, \quad if \quad |i \wedge j| = 0.$$

Here,  $|i \wedge j|$  represents the level of the highest common descendant  $i \wedge j$  of i and j (cf. Section 3.1). Note that correlations only disappears for leaves (i,j) such that  $n-|i \wedge j| \to \infty$ , that is, when the length of the path from i (or j) to the highest common descendant  $i \wedge j$  diverges as the height n of the tree increases. Unfortunately, correlations are non-negligible in all other cases even for  $n \to \infty$ . This imposes severe problems when handling formula (5.20) since products and expectations cannot be interchanged.

#### 6. Uniqueness of equilibria and quantitative ergodicity property

In this part we shall prove our main results, namely, Theorem 1.1 and Corollary 1.2. First, we will derive a local convergence result as  $n \to \infty$  of the profiles  $\{F_n\}_{n \in \mathbb{N}}$  solving (1.1) with generic initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  towards the Gaussian profile  $F_\alpha$  in (1.12). This weaker local analogue of Theorem 1.1 will be enough to derive the uniqueness of the solution  $(\lambda_\alpha, F_\alpha)$  of the eigenproblem (1.11) in Corollary 1.2. Second, we will extend our local result into our main global result in Theorem 1.1 by quantifying the relaxation of the growth rate  $\|F_n\|_{L^1(\mathbb{R}^d)}/\|F_{n-1}\|_{L^1(\mathbb{R}^d)}$  and the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  towards the unique solution  $(\lambda_\alpha, F_\alpha)$  of the eigenproblem (1.11). Our strategy will exploit the high-dimensional iterative formula (5.20) in Proposition 5.9.

6.1. Local convergence result. In this part we shall prove our local convergence result. To this end, we will show that the dependency of the expectation term in (5.20) on the variable x decays to zero as  $n \to \infty$  uniformly over compact sets. To simplify notation, we shall define the following operator.

Definition 6.1 (Expectation operators). Under the notation in Definitions 5.6 and 5.8, we define

$$\mathcal{E}_n[F](x) := \mathbb{E}\left[\prod_{j \in \mathbf{L}^n} F(\Phi_n^j(x; \mathbf{Y}_n))\right], \quad x \in \mathbb{R}^d,$$
(6.1)

for any  $n \in \mathbb{N}$  and any non-negative measurable function  $F : \mathbb{R}^d \to \mathbb{R}$ .

Note that  $\mathcal{E}_n[F](x)$  possibly takes extended values in  $(-\infty, +\infty]$  if the expectation does not exist. However, if F is a bounded function (and it will be often the case), we obtain that  $\mathcal{E}_n[F](x)$  always exists and is finite. The following technical lemma for the tail of the incomplete Gamma function will be required later to determine the asymptotic behavior of the above operator  $\mathcal{E}_n$ .

**Lemma 6.2.** Consider the (upper) incomplete Gamma function with parameter  $a \in \mathbb{R}^+_+$ , i.e.,

$$\Gamma(a,x) := \int_{x}^{+\infty} s^{a-1} e^{-s} ds, \quad x \in \mathbb{R}_{+}^{*}.$$
 (6.2)

Then, the following estimate holds true

$$\Gamma(a+1,x) \le 2e^{-x/2},$$
(6.3)

for every  $x, a \in \mathbb{R}_+^*$  verifying the constraints

$$x > e, \quad 2a < \frac{x}{\log(x)}. \tag{6.4}$$

*Proof.* Let us restate the incomplete Gamma function as follows

$$\Gamma(a+1,x) = \int_{x}^{+\infty} \exp(-\psi_a(s)s) \, ds,$$

where  $\psi_a(s)$  is the function

$$\psi_a(s) := 1 - a \frac{\log(s)}{s}, \quad s > 0.$$

Notice that  $\psi_a$  is monotonically increasing for  $s \in [e, \infty)$ . Hence, we obtain that

$$\psi_a(s) \ge \psi_a(x) \ge \frac{1}{2},$$

for every  $s \ge x$ , where we have used (6.4) on each of the above inequalities. Consequently, we get

$$\Gamma(a+1,x) \le \int_{r}^{+\infty} e^{-\frac{s}{2}} ds,$$

and this ends the proof.

Therefore, we obtain the following weak dependency of  $\mathcal{E}_n[F]$  on the variable x as  $n \to \infty$ .

**Lemma 6.3** (Asymptotics of  $\mathcal{E}_n$ ). Assume that  $\alpha \in \mathbb{R}_+^*$  and consider any  $F \in C^1(\mathbb{R}^d)$  such that

$$0 < F(x) \le C, |\nabla \log F(x)| \le D(1+|x|),$$
 (6.5)

for every  $x \in \mathbb{R}^d$  and appropriate  $C, D \in \mathbb{R}_+^*$ . Then, for every  $\varepsilon \in \mathbb{R}_+^*$ , there exists a large enough constant  $C_{\varepsilon} \in \mathbb{R}_+^*$  such that the following property holds true

$$|\nabla \log \mathcal{E}_n[F](x)| \le C_{\varepsilon} \left( (2 \, \mathbf{k}_{\alpha} + \varepsilon)^n + (\mathbf{r}_{\alpha} + \varepsilon)^n |x| + e^{-\frac{(2+\varepsilon)^n}{\varepsilon}} e^{C_{\varepsilon} ((2 \, \mathbf{r}_{\alpha} + \varepsilon)^n |x|^2)} \right), \tag{6.6}$$

for every  $x \in \mathbb{R}^d$  and any  $n \in \mathbb{N}$ , where  $\mathcal{E}_n$  is the operator in Definition 6.1.

*Proof.* Before proving the main result, note that the hypothesis (6.5) guarantee that F is bounded below by a Gaussian profile. Indeed, by the fundamental theorem of calculus we obtain the following relation

$$\log F(x) = \log F(0) + \int_0^1 \nabla(\log F)(\theta x) \cdot x \, d\theta,$$

for any  $x \in \mathbb{R}^d$ . Using the linear growth assumption in (6.5) for the log-derivative  $\nabla \log F$  we obtain that

$$\log F(x) \ge \log F(0) - D|x| \int_0^1 (1 + \theta|x|) \, d\theta \ge \log F(0) - \frac{D}{2} - D|x|^2,$$

for any  $x \in \mathbb{R}^d$ , where in the last inequality we have integrated with respect to  $\theta$  and we have used Young's inequality to interpolate |x| by  $|x|^2$ . Therefore, we obtain the following Gaussian lower bound

$$F(x) \ge ce^{-\beta|x|^2},\tag{6.7}$$

for each  $x \in \mathbb{R}^d$  and some  $c, \beta \in \mathbb{R}^*_+$  depending only on F(0) and D.

For simplicity of notation we define

$$E_n := \mathcal{E}_n[F],$$

for any  $n \in \mathbb{N}$ . In the sequel, we shall study the behavior of  $\nabla \log E_n$  as  $n \to \infty$ .

• Step 1: Control of  $\sum_{j \in L^n} |\Phi_n^j(x; \mathbf{y}_n)|$ .

By definition (5.15) of the lineage map  $\Phi_n^j$ , we obtain that

$$\sum_{j \in \mathsf{L}^n} |\Phi_n^j(x; \mathbf{y}_n)| = \sum_{j \in \mathsf{L}^n} \left| \kappa_n x + \sum_{m=0}^{n-1} \kappa_m y_{\mathsf{c}^m(j)} \right| \le 2^n \kappa_n |x| + \sum_{m=0}^{n-1} \kappa_m 2^m \sum_{i \in \mathsf{L}_{n-m}^n} |y_i|,$$

where we are using the tree-indexed notation  $\mathbf{y}_n = (y_i)_{i \in \mathsf{T}^n_*} \in \mathbb{R}^{2(2^n-1)d}$  in Remark 3.1 and  $\mathbf{c} : \mathsf{T}^n_* \longrightarrow \widehat{\mathsf{T}}^n$  is defined in Section 3.1. Note that in the last inequality we have used that each index of the level  $\mathsf{L}^n_{n-m}$  in the tree contributes to the amount of  $2^m$  leaves. Then, by the Cauchy–Schwarz inequality we obtain

$$\sum_{j \in \mathsf{L}^n} \left| \Phi_n^j(x; \mathbf{y}_n) \right| \le 2^n \kappa_n |x| + \left( \sum_{m=0}^{n-1} 2^{n-m} (2^m \kappa_m)^2 \right)^{1/2} \|\mathbf{y}_n\| \le 2^n \kappa_n |x| + \sqrt{2} \, 2^{n/2} \|\mathbf{y}_n\|, \tag{6.8}$$

for any  $\mathbf{y}_n \in \mathbb{R}^{2(2^n-1)d}$ , where in the last part we have used that

$$\sum_{m=0}^{n-1} 2^{n-m} (2^m \kappa_m)^2 \le 2^n \sum_{m=0}^{n-1} \frac{1}{2^m (1+\alpha)^{2m}} \le 2^n \sum_{m=0}^{\infty} \frac{1}{2^m (1+\alpha)^{2m}} \le 2^{n+1}.$$

#### • STEP 2: Derivative of $\log E_n$ .

Taking derivatives on (6.1), using the second property in (6.5) along with (6.8) entails

$$|\nabla \log E_{n}(x)| \leq \kappa_{n} \frac{\mathbb{E}\left[\left(\sum_{j \in \mathsf{L}^{n}} |\nabla \log F(\Phi_{n}^{j}(x; \mathbf{Y}_{n}))|\right) \prod_{j \in \mathsf{L}^{n}} F(\Phi_{n}^{j}(x; \mathbf{Y}_{n}))\right]}{\mathbb{E}\left[\prod_{j \in \mathsf{L}^{n}} F(\Phi_{n}^{j}(x; \mathbf{Y}_{n}))\right]}$$

$$\lesssim \kappa_{n} \frac{\mathbb{E}\left[\left(\sum_{j \in \mathsf{L}^{n}} (1 + |\Phi_{n}^{j}(x; \mathbf{Y}_{n})|)\right) \prod_{j \in \mathsf{L}^{n}} F(\Phi_{n}^{j}(x; \mathbf{Y}_{n}))\right]}{\mathbb{E}\left[\prod_{j \in \mathsf{L}^{n}} F(\Phi_{n}^{j}(x; \mathbf{Y}_{n}))\right]}$$

$$\lesssim 2^{n} \kappa_{n} + 2^{n} \kappa_{n}^{2} |x| + 2^{n/2} \kappa_{n} R_{n}(x),$$

$$(6.9)$$

for every  $x \in \mathbb{R}^d$ , where the remainder  $R_n$  reads

$$R_n(x) := \frac{\mathbb{E}\left[\|\mathbf{Y}_n\| \prod_{j \in \mathsf{L}^n} F(\Phi_n^j(x; \mathbf{Y}_n))\right]}{\mathbb{E}\left[\prod_{j \in \mathsf{L}^n} F(\Phi_n^j(x; \mathbf{Y}_n))\right]}, \quad x \in \mathbb{R}^d.$$

This estimation provides a first insight about the ergodicity property. Indeed the correlations, as measured by  $\kappa_n$ , decreases faster than  $2^{-n}$  (cf. Lemma 5.3), meaning that the contribution  $2^n\kappa_n$  decays fast to zero. It is also the case of the other contributions, locally in x, but the last one, namely  $2^{n/2}\kappa_n R_n(x)$ , requires more work. Consider any  $\sqrt{2} < q < \frac{1}{\sqrt{2} \mathbf{k}_{\alpha}}$ , which exists because  $2 \mathbf{k}_{\alpha} < 1$  for  $\alpha \in \mathbb{R}_+^*$ , and split  $R_n(x) = R_n^1(x) + R_n^2(x)$  in terms of the functions

$$R_n^1(x) := \frac{1}{Z_n(x)} \int_{\|\mathbf{y}_n\| \le q^n} \|\mathbf{y}_n\| \prod_{j \in \mathbf{L}^n} F(\Phi_n^j(x; \mathbf{y}_n)) e^{-Q_n(\mathbf{y}_n)} \, d\mathbf{y}_n, \tag{6.10}$$

$$R_n^2(x) := \frac{1}{Z_n(x)} \int_{\|\mathbf{y}_n\| > q^n} \|\mathbf{y}_n\| \prod_{j \in \mathsf{L}^n} F(\Phi_n^j(x; \mathbf{y}_n)) e^{-Q_n(\mathbf{y}_n)} d\mathbf{y}_n, \tag{6.11}$$

and the normalization factor

$$Z_n(x) := \int_{\mathbb{R}^{2(2^n - 1)d}} \prod_{j \in \mathbb{L}^n} F(\Phi_n^j(x; \mathbf{y}_n)) e^{-Q_n(\mathbf{y}_n)} \, d\mathbf{y}_n.$$
 (6.12)

#### • Step 3: Lower bound for $Z_n$ .

Using the above lower bound by a Gaussian in (6.7) along with the upper bound for  $Q_n$  in Lemma 5.5

implies the following estimate of  $Z_n$  in (6.12):

$$\begin{split} Z_n(x) &\geq c^{2^n} \int_{\mathbb{R}^{2(2^n-1)d}} \exp\left(-\beta \sum_{j \in \mathbb{L}^n} |\Phi_n^j(x; \mathbf{y}_n)|^2\right) \exp(-\beta_{\max} \|\mathbf{y}_n\|^2) \, d\mathbf{y}_n \\ &\geq c^{2^n} e^{-2^{2n+1}\kappa_n^2 \beta \|x\|^2} \int_{\mathbb{R}^{2(2^n-1)d}} \exp\left(-(2^{n+2}\beta + \beta_{\max}) \|\mathbf{y}_n\|^2\right) d\mathbf{y}_n \\ &= c^{2^n} e^{-2^{2n+1}\kappa_n^2 \beta |x|^2} \left|\mathbb{S}^{2(2^n-1)d-1}\right| \int_0^{+\infty} r^{2(2^n-1)d-1} \exp\left(-(2^{n+2}\beta + \beta_{\max})r^2\right) \, dr \\ &= \frac{c^{2^n} e^{-2^{2n+1}\kappa_n^2 \beta |x|^2}}{2(2^{n+2}\beta + \beta_{\max})(2^{n-1)d}} \left|\mathbb{S}^{2(2^n-1)d-1}\right| \int_0^{+\infty} s^{(2^n-1)d-1} e^{-s} \, ds \\ &= \frac{c^{2^n} e^{-2^{2n+1}\kappa_n^2 \beta |x|^2}}{2(2^{n+2}\beta + \beta_{\max})(2^{n-1)d}} \left|\mathbb{S}^{2(2^n-1)d-1}\right| \Gamma((2^n-1)d), \end{split}$$

where in the last step we have used the Gamma function, that is defined by

$$\Gamma(a) := \int_0^{+\infty} s^{a-1} e^{-s} \, ds, \quad a > 0.$$

Recall that area of the hypersphere is  $|\mathbb{S}^{d-1}| = \frac{2\pi^{d/2}}{\Gamma(\frac{d}{2})}$ . Then, we obtain

$$Z_n(x) \ge \frac{c^{2^n} \pi^{(2^n - 1)d}}{(2^{n+2}\beta + \beta_{\max})^{(2^n - 1)d}} e^{-2^{2n+1} \kappa_n^2 \beta |x|^2},$$
(6.13)

for every  $x \in \mathbb{R}^d$ .

• STEP 4: Upper bound for  $R_n^1$  and  $R_n^2$ . On the one hand, notice that

$$R_n^1(x) \le q^n, \tag{6.14}$$

for all  $x \in \mathbb{R}^d$  by definition of  $R_n^1$  in (6.10). On the other hand, using the lower bound of  $Q_n$  in Lemma 5.5 we can estimate  $R_n^2$  in (6.11) by

$$R_{n}^{2}(x) \leq \frac{C^{2^{n}}}{q^{n}Z_{n}(x)} \int_{\|\mathbf{y}_{n}\| \geq q^{n}} \|\mathbf{y}_{n}\|^{2} e^{-\beta_{\min}\|\mathbf{y}_{n}\|^{2}} d\mathbf{y}_{n}$$

$$= \frac{C^{2^{n}}}{q^{n}Z_{n}(x)} \left| \mathbb{S}^{2(2^{n}-1)d-1} \right| \int_{q^{n}}^{+\infty} r^{2(2^{n}-1)d+1} e^{-\beta_{\min}r^{2}} dr$$

$$= \frac{C^{2^{n}}}{q^{n}Z_{n}(x)} \left| \mathbb{S}^{2(2^{n}-1)d-1} \right| \frac{1}{2\beta_{\min}^{(2^{n}-1)d+1}} \int_{\beta_{\min}q^{2n}}^{+\infty} s^{(2^{n}-1)d} e^{-s} ds$$

$$= \frac{C^{2^{n}}\pi^{(2^{n}-1)d}}{q^{n}\beta_{\min}^{(2^{n}-1)d+1}Z_{n}(x)} \frac{\Gamma\left((2^{n}-1)d+1, \beta_{\min}q^{2n}\right)}{\Gamma((2^{n}-1)d)}, \tag{6.15}$$

for all  $x \in \mathbb{R}^d$ , where we have used again the incomplete Gamma function (6.2) in Lemma 6.2. Putting the lower estimate (6.13) for  $Z_n$  into (6.15) yields

$$R_n^2(x) \lesssim e^{2^{2n+1}\kappa_n^2\beta|x|^2} \frac{C^{2^n}(2^{n+2}\beta + \beta_{\max})^{(2^n-1)d}}{q^n c^{2^n}\beta_{\min}^{(2^n-1)d}} \frac{\Gamma\left((2^n-1)d + 1, \beta_{\min}q^{2n}\right)}{\Gamma((2^n-1)d)},\tag{6.16}$$

for all  $x \in \mathbb{R}^d$ . Now, take  $a = (2^n - 1)d$  and  $x = \beta_{\min}q^{2n}$  and notice that the constraints (6.4) in Lemma 6.2 hold true for large enough  $n \in \mathbb{N}$  as long as we choose  $q > \sqrt{2}$ . Hence, we obtain the estimate

$$\Gamma\left((2^n - 1)d + 1, \beta_{\min}q^{2n}\right) \le 2e^{-\frac{\beta_{\min}}{2}q^{2n}},$$

for all  $n \ge n_* = n_*(\beta_{\min}, d, q)$ . Using Stirling's formula for the remaining Gamma function in (6.16), we conclude the estimate

$$R_n^2(x) \lesssim e^{2^{2n+1}\kappa_n^2\beta|x|^2} \frac{C^{2^n}(2^{n+2}\beta + \beta_{\max})^{(2^n-1)d}}{q^n c^{2^n}\beta_{\min}^{(2^n-1)d}} \frac{e^{-\frac{\beta_{\min}}{2}}q^{2^n}}{(2^n-1)^{1/2}\left(\frac{(2^n-1)d}{e}\right)^{(2^n-1)d}},\tag{6.17}$$

for every  $x \in \mathbb{R}^d$  and each  $n \geq n_*$ .

• Step 5: Final conclusion.

Let us put the preceding estimates (6.14) and (6.17) into (6.9) to achieve

$$|\nabla \log E_n(x)| \lesssim 2^n \kappa_n + 2^{n/2} q^n \kappa_n + 2^n \kappa_n^2 |x| + \frac{M^{2^n}}{e^{\frac{\beta_{\min}}{2} q^{2^n}}} e^{2^{2n+1} \kappa_n^2 \beta |x|^2},$$
(6.18)

for every  $x \in \mathbb{R}^d$  and each  $n \ge n_*$ , where  $M = M(C, c, \beta, \beta_{\min}, \beta_{\max}, d, q)$  is a universal constant that we have introduced to absorb all the double exponential factors of the form  $a^{2^n}$  with a > 0 in the fourth term along with any further lower order factor. On the one hand, choosing  $q > \sqrt{2}$  sufficiently close to  $\sqrt{2}$  and using (5.11) in Lemma 5.3 we obtain the following asymptotics for the factors in the first three terms of (6.18):

$$\begin{split} 2^n \kappa_n &\leq C_{\varepsilon} (2 \, \boldsymbol{k}_{\alpha} + \varepsilon)^n, \\ 2^{n/2} q^n \kappa_n &\leq C_{\varepsilon} (2 \, \boldsymbol{k}_{\alpha} + \varepsilon)^n, \\ 2^n \kappa_n^2 &\leq C_{\varepsilon} (2 (\boldsymbol{k}_{\alpha} + \varepsilon)^2)^n \leq C_{\varepsilon} (\boldsymbol{r}_{\alpha} + \varepsilon)^n, \end{split}$$

for arbitrarily small  $\varepsilon \in \mathbb{R}_+^*$  and sufficiently large  $C_{\varepsilon} \in \mathbb{R}_+^*$ . Note that in the last identity we have used the relation  $\mathbf{r}_{\alpha} = 2 \mathbf{k}_{\alpha}^2$ . On the other hand, since q has been taken larger than  $\sqrt{2}$  and arbitrarily close to  $\sqrt{2}$ , then the double exponential in the denominator of the last term in (6.18) kills the double exponential on the numerator independently on the explicit value of M. Therefore, we obtain

$$\frac{M^{2^n}}{e^{\frac{\beta_{\min}}{2}q^{2n}}} \le C_{\varepsilon} e^{-\frac{(2+\varepsilon)^n}{\varepsilon}},$$

by appropriately increasing  $C_{\varepsilon}$  if necessary. Similarly, we infer the following control for the remaining x-dependent exponential factor in the fourth term of (6.18)

$$e^{2^{2n+1}\kappa_n^2\beta|x|^2} < e^{C_{\varepsilon}(4(\boldsymbol{k}_{\alpha}+\varepsilon)^2)^n|x|^2} < e^{C_{\varepsilon}(2\boldsymbol{r}_{\alpha}+\varepsilon)^n|x|^2},$$

which yields the announced estimate (6.6).

Remark 6.4 (Local convergence). Under the assumptions in Lemma 6.3, we obtain that

$$\lim_{n \to \infty} \sup_{|x| \le R} |\nabla \log \mathcal{E}_n[F](x)| = 0,$$

for any radius  $R \in \mathbb{R}^*_{\perp}$ .

Therefore, we arrive at the following local version of our main Theorem 1.1.

**Corollary 6.5** (Local convergence result). Assume that  $\alpha \in \mathbb{R}_+^*$ , set any initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  and consider the solution  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1). Then,

$$\lim_{n \to \infty} \sup_{|x| \le R} \left| \frac{F_n(x)}{F_n(0)} - \frac{\boldsymbol{F}_{\alpha}(x)}{\boldsymbol{F}_{\alpha}(0)} \right| = 0, \tag{6.19}$$

for every  $R \in \mathbb{R}_+^*$ , where  $\mathbf{F}_{\alpha}$  is the Gaussian profile (1.12).

*Proof.* By formula (5.20) in Proposition (5.9) and Definition 6.1 we obtain the following relation

$$\nabla \log F_n(x) = -(1 + \alpha - k_n)x + \nabla \log \mathcal{E}_n[\bar{F}_0](x),$$

for any  $x \in \mathbb{R}^d$  and every  $n \in \mathbb{N}$ . Using the mean value theorem we then achieve

$$\log \frac{F_n(x)}{F_n(0)} - \log \frac{\mathbf{F}_{\alpha}(x)}{\mathbf{F}_{\alpha}(0)} = \int_0^1 (\nabla \log F_n)(\theta x) \cdot x \, d\theta + \frac{1}{2\sigma_{\alpha}^2} |x|^2$$

$$= -\frac{1 + \alpha - k_n}{2} |x| + \int_0^1 \nabla \log \mathcal{E}_n[\bar{F}_0](\theta x) \cdot x \, d\theta + \frac{1}{2\sigma_{\alpha}^2} |x|^2$$

$$= \frac{k_n - \mathbf{k}_{\alpha}}{2} |x|^2 + \int_0^1 \nabla \log \mathcal{E}_n[\bar{F}_0](\theta x) \cdot x \, d\theta,$$
(6.20)

where we have used the relation  $1/\sigma_{\alpha}^2 = 1 + \alpha - \mathbf{k}_{\alpha}$ . On the one hand, the first term in the right hand side of (6.20) converges to zero thanks to Lemma 5.3. On the other hand, for the second term we shall apply Lemma 6.3. To do so we need to guarantee that  $\bar{F}_0$  verifies the hypothesis (6.5) of such a lemma. Although this is not true for generic initial data, note that after two iterations  $\bar{F}_2$  already satisfies (6.5)<sub>2</sub> thanks to Corollary 4.2. Regarding (6.5)<sub>1</sub> note that, unfortunately, only two iterations are not enough to guarantee such a crucial boundedness hypothesis in general. However, again by Corollary 4.2 recall that

$$\bar{F}_m(x) \le C_m e^{\frac{1}{2} \left(\alpha + \frac{1}{2} - \frac{1}{\bar{\sigma}_m^2}\right)|x|^2},$$

for appropriate  $C_m \in \mathbb{R}_+^*$  and the sequence of variances  $\{\overline{\sigma}_m^2\}_{m \in \mathbb{N}}$  verifying the recurrence (4.5) with initial datum  $\overline{\sigma}_1^2 = 1/\alpha$ . Note then that

$$\alpha + \frac{1}{2} - \frac{1}{\overline{\sigma}_{m}^{2}} = \left(\alpha + \frac{1}{2} - \frac{1}{\sigma_{\alpha}^{2}}\right) + \left(\frac{1}{\sigma_{\alpha}^{2}} - \frac{1}{\overline{\sigma}_{m}^{2}}\right) \leq \frac{\sigma_{\alpha}^{2} - 2}{4} + C_{v} r_{\alpha}^{m},$$

for any  $m \in \mathbb{N}$ , where we have used Lemma 3.8 in the second term. Since  $\sigma_{\alpha}^2 < 2$  because  $\alpha \in \mathbb{R}_+^*$  (recall Figure 2b) and  $\mathbf{r}_{\alpha} \in (0,1)$  we then obtain that the above coefficient becomes negative if  $m \geq n_{\alpha}$  for a sufficiently large  $n_{\alpha} \geq 2$  (depending only on  $\alpha$ ). This justifies that  $\bar{F}_{n_{\alpha}}$  satisfies both conditions in (6.5). Hence, we shall actually assume that  $\bar{F}_0$  already satisfies (6.5) without loss of generality. Therefore, we can apply Remark 6.4 with  $F = \bar{F}_0$  and obtain that  $\nabla \log \mathcal{E}_n[\bar{F}_0]$  converges to zero uniformly over compact sets. Consequently, taking limits in (6.20) entails

$$\lim_{n \to \infty} \sup_{|x| \le R} \left| \log \frac{F_n(x)}{F_n(0)} - \log \frac{\boldsymbol{F}_\alpha(x)}{\boldsymbol{F}_\alpha(0)} \right| = 0,$$

for any  $R \in \mathbb{R}_+^*$ , thus proving (6.19).

Proof of Corollary 1.2. Let  $(\lambda, F)$  be any solution of the eigenproblem (1.11). Then, the ansatz (1.10) defines a solution  $F_n = \lambda^n F$  of the time-discrete problem (1.1). By Corollary 6.5 we obtain that

$$\sup_{|x| < R} \left| \frac{F(x)}{F(0)} - \frac{\mathbf{F}_{\alpha}(x)}{\mathbf{F}_{\alpha}(0)} \right| = 0,$$

for any  $R \in \mathbb{R}_+^*$ . Hence,  $F/F(0) \equiv \mathbf{F}_{\alpha} / \mathbf{F}_{\alpha}(0)$ . Since both F and  $\mathbf{F}_{\alpha}$  are normalized distributions, we conclude that  $F \equiv \mathbf{F}_{\alpha}$  and thus  $\lambda = \lambda_{\alpha}$ .

Whilst the above local convergence result in Corollary 6.5 is enough to identify asymptotically the profile  $F_{\alpha}$ , a global result with quantitative convergence rates is still missing. In addition, the identification of the associated eigenvalue  $\lambda_{\alpha}$  as a limiting quantity is not apparent via this method. In the following section, we give an answer to both questions by better exploiting the previous fundamental Lemma 6.3 and using the propagation of quadratic and exponential moments in Section 4.2.

6.2. Global convergence result. Before proving our main result, we require a careful control of the ratios  $F_n(0)/\|F_n\|_{L^1(\mathbb{R}^d)}$  as  $n \to \infty$ . Indeed, if  $\{F_n\}_{n \in \mathbb{N}}$  is the solution (3.4) of the time-discrete problem (1.1) issued at a Gaussian initial data  $F_0$  like in Proposition 3.7 we obtain explicitly

$$\frac{F_n(0)}{\|F_n\|_{L^1(\mathbb{R}^d)}} = G_{\mu_n,\sigma_n^2}(0) \to \pmb{F}_\sigma(0) = \frac{1}{(2\pi\,\pmb{\sigma}_\alpha^2)^{d/2}},$$

as  $n \to \infty$ , where  $\mu_n$  and  $\sigma_n^2$  are determined by the recursive relations (3.5) and we have used that  $\mu_n \to 0$  and  $\sigma_n^2 \to \sigma_\alpha^2$  according to Lemma 3.8. In the following lemma, we provide a similar quantitative result with convergence rates for generic initial data  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$ .

**Lemma 6.6.** Assume that  $\alpha \in \mathbb{R}_+^*$ , set any initial datum  $F_0 \in \mathcal{M}_+(\mathbb{R}^d)$  and consider the solution  $\{F_n\}_{n\in\mathbb{N}}$  to the time-discrete problem (1.1). For any  $\varepsilon \in \mathbb{R}_+^*$  there exists  $C_\varepsilon \in \mathbb{R}_+^*$  sufficiently large with

$$\left| \frac{F_n(0)}{\|F_n\|_{L^1(\mathbb{R}^d)}} - \frac{1}{(2\pi \boldsymbol{\sigma}_{\alpha}^2)^{d/2}} \right| \le C_{\varepsilon} ((2\boldsymbol{k}_{\alpha})^{2/3} + \varepsilon)^n, \tag{6.21}$$

for any  $n \in \mathbb{N}$ .

*Proof.* Let us fix any  $\varepsilon \in \mathbb{R}_+^*$  arbitrarily small so that  $(2 \mathbf{k}_{\alpha})^{2/3} + \varepsilon < 1$ . Again, this is true because  $2 \mathbf{k}_{\alpha} < 1$  for  $\alpha \in \mathbb{R}_+^*$ . For simplicity of notation, we shall define the following sequence of coefficients:

$$c_n := \frac{F_n(0)}{\|F_n\|_{L^1(\mathbb{R}^d)}}, \quad n \in \mathbb{N}.$$

Our goal then reduces to studying the asymptotic behavior of  $\{c_n\}_{n\in\mathbb{N}}$  and obtaining quantitative convergence rates. By formula (5.20) in Proposition 5.9 for the recursion, we obtain

$$\frac{1}{c_n} = \frac{\int_{\mathbb{R}^d} e^{-\frac{1+\alpha-k_n}{2}|x|^2} \mathcal{E}_n[\bar{F}_0](x) dx}{\mathcal{E}_n[\bar{F}_0](0)} = \int_{|x| \le R_n} e^{-\frac{1+\alpha-k_n}{2}|x|^2} \frac{\mathcal{E}_n[\bar{F}_0](x)}{\mathcal{E}_n[\bar{F}_0](0)} dx + \frac{1}{c_n} \int_{|x| > R_n} \frac{F_n(x)}{\|F_n\|_{L^1(\mathbb{R}^d)}} dx,$$

where we set the sequence of radius  $\{R_n\}_{n\in\mathbb{N}}$  as follows

$$R_n = \frac{1}{(2 \mathbf{k}_{\alpha})^{n/3}}, \quad n \in \mathbb{N}.$$
(6.22)

By solving the implicit equation on  $c_n$  we infer

$$c_n = \frac{1 - \tilde{E}_n}{(2\pi \,\sigma_\alpha^2)^{d/2} + E_n},\tag{6.23}$$

for every  $n \in \mathbb{N}$ , where the error terms  $\tilde{E}_n$  and  $E_n = E_{n,1} + E_{n,2} + E_{n,3}$  take the form

$$\tilde{E}_{n} := \int_{|x| > R_{n}} \frac{F_{n}(x)}{\|F_{n}\|_{L^{1}(\mathbb{R}^{d})}} dx, 
E_{n,1} := \left(\frac{2\pi}{1 + \alpha - k_{n}}\right)^{d/2} - \left(\frac{2\pi}{1 + \alpha - k_{\alpha}}\right)^{d/2}, 
E_{n,2} := \int_{|x| > R_{n}} e^{-\frac{1 + \alpha - k_{n}}{2}|x|^{2}} dx, 
E_{n,3} := \int_{|x| \le R_{n}} e^{-\frac{1 + \alpha - k_{n}}{2}|x|^{2}} \left(\frac{\mathcal{E}_{n}[\bar{F}_{0}](x)}{\mathcal{E}_{n}[\bar{F}_{0}](0)} - 1\right) dx.$$
(6.24)

On the one hand, the propagation of quadratic moments in Corollary 4.6 entails

$$\tilde{E}_n \le \frac{1}{R_n^2} \int_{\mathbb{R}^d} |x|^2 \frac{F_n(x)}{\|F_n\|_{L^1(\mathbb{R}^d)}} \lesssim \frac{1}{R_n^2} \lesssim (2 \, \mathbf{k}_\alpha)^{2n/3}.$$
 (6.25)

On the other hand, by Lemma 5.3 we readily infer that  $E_{n,1}$  in (6.24) can be estimated by

$$|E_{n,1}| \lesssim k_n - \mathbf{k}_\alpha \lesssim \mathbf{r}_\alpha^n$$
. (6.26)

By direct calculation we obtain

$$E_{n,2} = |\mathbb{S}^{d-1}| \int_{R_n}^{+\infty} r^{d-1} e^{-\frac{1+\alpha-k_n}{2}r^2} \lesssim \Gamma\left(\frac{d}{2}, \frac{1+\alpha-k_n}{2}R_n^2\right) \lesssim e^{-\frac{1+\alpha-k_n}{4(2k_\alpha)^{2n/3}}},\tag{6.27}$$

where  $\Gamma(a,x)$  denotes the incomplete Gamma function (6.2) and we have used estimate (6.3) in Lemma 6.2 as the constraint (6.4) is trivially satisfied by our choice (6.22) since  $R_n \to \infty$ . Finally, recall that we can assume that  $\bar{F}_0$  satisfies the hypothesis (6.5) in Lemma 6.3 without loss of generality (*cf.* proof of Corollary 6.5). Then, given any  $\varepsilon' \in \mathbb{R}_+^*$  there exists a large enough  $C_{\varepsilon'} \in \mathbb{R}_+^*$  so that

$$\begin{split} & \left| \int_{0}^{1} (\nabla \log \mathcal{E}_{n}) [\bar{F}_{0}] (\theta x) \cdot x \, d\theta \right| \\ & \lesssim (2 \, \boldsymbol{k}_{\alpha} + \varepsilon')^{n} R_{n} + (\boldsymbol{r}_{\alpha} + \varepsilon')^{n} R_{n}^{2} + e^{-\frac{(2+\varepsilon')^{n}}{\varepsilon'}} e^{C_{\varepsilon'} (2 \, \boldsymbol{r}_{\alpha} + \varepsilon')^{n} R_{n}^{2}} R_{n} \\ & \lesssim \left( \frac{2 \, \boldsymbol{k}_{\alpha}}{(2 \, \boldsymbol{k}_{\alpha})^{1/3}} + \varepsilon \right)^{n} + \left( \frac{\boldsymbol{r}_{\alpha}}{(2 \, \boldsymbol{k}_{\alpha})^{2/3}} + \varepsilon \right)^{n} + e^{-\frac{(2+\varepsilon')^{n}}{\varepsilon'}} e^{C_{\varepsilon'} \left( \frac{2 \, \boldsymbol{r}_{\alpha}}{(2 \, \boldsymbol{k}_{\alpha})^{2/3}} + \varepsilon \right)^{n}} \frac{1}{(2 \, \boldsymbol{k}_{\alpha})^{n/3}} \\ & \lesssim ((2 \, \boldsymbol{k}_{\alpha})^{2/3} + \varepsilon)^{n}, \end{split}$$

where in the third line we have taken  $\varepsilon'$  small enough compared with  $\varepsilon$  and in the last line we have used the relation  $\mathbf{r}_{\alpha} = 2 \mathbf{k}_{\alpha}^2$ , which guarantees that all the contributions in the third line are dominated by the first term  $((2 \mathbf{k}_{\alpha})^{2/3} + \varepsilon)^n$ . Therefore, the mean value theorem implies

$$\left| \frac{\mathcal{E}_n[\bar{F}_0](x)}{\mathcal{E}_n[\bar{F}_0](0)} - 1 \right| = \left| \exp\left( \int_0^1 (\nabla \log \mathcal{E}_n[\bar{F}_0])(\theta x) \cdot x \, d\theta \right) - 1 \right| \lesssim ((2 \, \boldsymbol{k}_\alpha)^{2/3} + \varepsilon)^n.$$

Hence,  $E_n^3$  in (6.24) can be controlled by

$$|E_n^3| = \int_{|x| \le R_n} e^{-\frac{1 + \alpha - k_n}{2}|x|^2} \left| \frac{\mathcal{E}_n[\bar{F}_0](x)}{\mathcal{E}_n[\bar{F}_0](0)} - 1 \right| dx \lesssim ((2 \mathbf{k}_\alpha)^{2/3} + \varepsilon)^n.$$
 (6.28)

Putting (6.26), (6.27) and (6.28) together yields

$$|E_n| \lesssim ((2 \mathbf{k}_\alpha)^{2/3} + \varepsilon)^n. \tag{6.29}$$

Thereby, using the mean value theorem on (6.23) along with the bounds (6.25) and (6.29) concludes that

$$\left| c_n - \frac{1}{(2\pi \sigma_{\alpha}^2)^{d/2}} \right| \lesssim \tilde{E}_n + |E_n| \lesssim ((2 \mathbf{k}_{\alpha})^{2/3} + \varepsilon)^n.$$

We are now ready to prove our main result.

Proof of Theorem 1.1.

• STEP 1: Convergence of the profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$ .

By definition of the relative entropy we obtain the decomposition:

$$\mathcal{D}_{\mathrm{KL}}\left(\frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} \| \mathbf{F}_{\alpha}\right) = D_{n,1} + D_{n,2},\tag{6.30}$$

where each factor takes the form

$$D_{n,1} := \int_{\mathbb{R}^d} \left( \log \frac{F_n(x)}{F_n(0)} - \log \frac{F_{\alpha}(x)}{F_{\alpha}(0)} \right) \frac{F_n(x)}{\|F_n\|_{L^1(\mathbb{R}^d)}} dx,$$

$$D_{n,2} := \log \frac{F_n(0)}{\|F_n\|_{L^1(\mathbb{R}^d)}} - \log F_{\alpha}(0).$$

On the one hand, for  $D_{n,1}$  we shall refine the argument in formula (6.20) in the proof of Corollary 6.5. Specifically, recall again that we can assume that  $\bar{F}_0$  satisfies the hypothesis (6.5) without loss of generality (cf. proof of Corollary 6.5). Then, applying Lemma 6.3 in order to control  $\nabla \log \mathcal{E}_n[\bar{F}_0]$  with fixed  $\varepsilon \in \mathbb{R}_+^*$  and sufficiently large  $C_{\varepsilon} \in \mathbb{R}_+^*$  yields

$$\left|\log \frac{F_n(x)}{F_n(0)} - \log \frac{F_\alpha(x)}{F_\alpha(0)}\right| \lesssim (2 \mathbf{k}_\alpha + \varepsilon)^n |x| + (\mathbf{r}_\alpha + \varepsilon)^n |x|^2 + e^{-\frac{(2+\varepsilon)^n}{\varepsilon}} e^{C_\varepsilon (2 \mathbf{r}_\alpha + \varepsilon)^n |x|^2} |x|,$$

for any  $x \in \mathbb{R}^d$ . Integrating the previous inequality against  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  and appropriately applying Jensen's inequality in the first term and the Cauchy–Schwarz inequality in the third term we obtain

$$|D_{n,1}| \lesssim (2 \, \mathbf{k}_{\alpha} + \varepsilon)^{n} \left( \int_{\mathbb{R}^{d}} |x|^{2} \frac{F_{n}(x)}{\|F_{n}\|_{L^{1}(\mathbb{R}^{d})}} \, dx \right)^{1/2}$$

$$+ (\mathbf{r}_{\alpha} + \varepsilon)^{n} \int_{\mathbb{R}^{d}} |x|^{2} \frac{F_{n}(x)}{\|F_{n}\|_{L^{1}(\mathbb{R}^{d})}} \, dx$$

$$+ e^{-\frac{(2+\varepsilon)^{n}}{\varepsilon}} \left( \int_{\mathbb{R}^{d}} e^{2C_{\varepsilon}(2 \, \mathbf{r}_{\alpha} + \varepsilon)^{n} |x|^{2}} \frac{F_{n}(x)}{\|F_{n}\|_{L^{1}(\mathbb{R}^{d})}} \, dx \right)^{1/2} \left( \int_{\mathbb{R}^{d}} |x|^{2} \frac{F_{n}(x)}{\|F_{n}\|_{L^{1}(\mathbb{R}^{d})}} \, dx \right)^{1/2}.$$

Take  $n_{\varepsilon} \in \mathbb{N}$  sufficiently large so that

$$\sup_{n>n_{\varepsilon}} 2C_{\varepsilon}(2\boldsymbol{r}_{\alpha}+\varepsilon)^{n} < \frac{\alpha}{2}$$

Then, Corollaries 4.6 and 4.8 imply that the above quadratic and exponential moments are uniformly bounded with respect to n for  $n \ge n_{\varepsilon}$ . Therefore,

$$|D_{n,1}| \lesssim (2 \, \mathbf{k}_{\alpha} + \varepsilon)^n + (\mathbf{r}_{\alpha} + \varepsilon)^n + e^{-\frac{(2+\varepsilon)^n}{\varepsilon}} \lesssim (2 \, \mathbf{k}_{\alpha} + \varepsilon)^n.$$
(6.31)

On the other hand, by the mean value theorem and Lemma 6.6 yields

$$|D_{n,2}| \lesssim \left| \frac{F_n(0)}{\|F_n\|_{L^1(\mathbb{R}^d)}} - \boldsymbol{F}_{\alpha}(0) \right| \lesssim ((2\,\boldsymbol{k}_{\alpha})^{2/3} + \varepsilon)^n. \tag{6.32}$$

Finally, putting (6.31) and (6.32) into (6.30) ends this part of the proof.

• STEP 2: Convergence of the growth rates  $||F_n||_{L^1(\mathbb{R}^d)}/||F_{n-1}||_{L^1(\mathbb{R}^d)}$ .

Let us fix any  $\varepsilon \in \mathbb{R}_+^*$  arbitrarily small so that  $(2 \mathbf{k}_{\alpha})^{2/3} + \varepsilon < 1$ . For simplicity of notation we shall define the following sequence of coefficients:

$$\lambda_n := \frac{\|F_n\|_{L^1(\mathbb{R}^d)}}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}}, \quad n \in \mathbb{N}.$$

Our goal then reduces to studying the asymptotic behavior of  $\{\lambda_n\}_{n\in\mathbb{N}}$  and obtaining quantitative convergence rates. By definition of operator  $\mathcal{T}$  in (1.2) we obtain that

$$\lambda_n = \int_{\mathbb{R}^d} e^{-m(x)} \int_{\mathbb{R}^{2d}} G\left(x - \frac{x_1 + x_2}{2}\right) \frac{F_{n-1}(x_1)}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} \frac{F_{n-1}(x_2)}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} dx_1 dx_2 dx.$$

By direct computation of the integral with respect to x we find that

$$\lambda_n = \int_{\mathbb{R}^{2d}} H(x_1, x_2) \frac{F_{n-1}(x_1)}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} \frac{F_{n-1}(x_2)}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} dx_1 dx_2, \tag{6.33}$$

where the function  $H(x_1, x_2)$  takes the form

$$H(x_1, x_2) := \frac{1}{(1+\alpha)^{d/2}} \exp\left(-\frac{\alpha}{(1+\alpha)} \left| \frac{x_1 + x_2}{2} \right|^2\right), \quad (x_1, x_2) \in \mathbb{R}^d.$$
 (6.34)

In addition, recalling the explicit form of the eigen-pair  $(\lambda_{\alpha}, F_{\alpha})$  in (1.12) implies

$$\boldsymbol{\lambda}_{\alpha} = \int_{\mathbb{R}^{2d}} H(x_1, x_2) \, \boldsymbol{F}_{\alpha}(x_1) \, \boldsymbol{F}_{\alpha}(x_2) \, dx_1 \, dx_2. \tag{6.35}$$

Therefore, taking the difference of (6.33) and (6.35) and noticing that H in (6.34) is bounded yields

$$|\lambda_n - \boldsymbol{\lambda}_{\alpha}| \lesssim \left\| \frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} \otimes \frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} - \boldsymbol{F}_{\alpha} \otimes \boldsymbol{F}_{\alpha} \right\|_{L^1(\mathbb{R}^{2d})} \lesssim \sqrt{\mathcal{D}_{\mathrm{KL}} \left( \frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} \right\| \boldsymbol{F}_{\alpha} \right)},$$

for any  $n \in \mathbb{N}$ , where in the last inequality we have used Pinsker's inequality. The result then follows from the above explicit convergence rates of the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  in STEP 1.

#### 7. Numerical experiments

For our numerical simulation, we have restricted to one-dimensional traits (i.e., d = 1) and we have considered a step function of the following form as initial datum:

$$F_0 = \frac{1}{Z} (30 \, \mathbb{1}_{[-7,-3]} + 20 \, \mathbb{1}_{[7.5,12.5]} + 50 \, \mathbb{1}_{[30,40]} + 30 \, \mathbb{1}_{[52,5,57,5]}), \tag{7.1}$$

where  $Z \in \mathbb{R}_+^*$  is a normalizing factor so that  $F_0 \in \mathcal{P}(\mathbb{R})$ . Let  $\{F_n\}_{n \in \mathbb{N}}$  be the solution to the timediscrete problem (1.1) starting at  $F_0$ . Our numerical simulation is performed with PYTHON on the finite computation domain [-15,60] for the trait variable x, which contains the support of the previous initial datum  $F_0$  and all the mass of each  $F_n$  (except for a negligible Gaussian tail). In our simulation, we set  $\Delta x = 0.001$  as the step for the discretization of our computational domain. In particular, with such a step we compute numerical integrals with respect to x according to the left rectangle rule.

In the following, we explore numerically two different scenarios: weak selection and strong selection. Specifically, we obtain numerical approximations for the profiles  $F_n$  in both regimes and we illustrate numerically the convergence of the growth rates  $||F_n||_{L^1(\mathbb{R}^d)}/||F_{n-1}||_{L^1(\mathbb{R}^d)}$  towards the eigenvalue  $\lambda_{\alpha}$ , and the relaxation of the normalized profiles  $F_n/||F_n||_{L^1(\mathbb{R}^d)}$  towards the eigenfunction  $F_{\alpha}$ . As a consequence, we derive numerical approximations of the convergence rates to be compared with the theoretical results in this paper. More specifically, as discussed previously in Remark 3.10 (see also Figure 4), we note that the theoretical convergence rates in Theorem 1.1 are unfortunately not sharp. Indeed, we actually attain numerically the same convergence rates as in Corollary 3.9, which were sharp for Gaussian initial data.

7.1. Weak selection. In this part, we consider a small value  $\alpha = 0.015$ . This leads to the following numerical values of the features of the equilibrium:

$$\lambda_{\alpha} \approx 0.9857, \quad \sigma_{\alpha}^2 \approx 1.8897.$$

characterizing the eigenpair  $(\lambda_{\alpha}, \mathbf{F}_{\alpha})$ , according to (1.12) and (1.13). Note that since the selection parameter has been taken very small, then the eigenvalue and the variance of the eigenfunction are close to those at linkage equilibrium, namely,  $\lambda_0 = 1$  and  $\sigma_0^2 = 2$  (see Remark 3.5 and Figure 2).

In Figure 6 we observe the relaxation of the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  towards the eigenfunction  $F_\alpha$  along the time iterations n=0,1,2,3,4,7,150. We remark on the strong contraction of the variance during the first few iterations leading to a well-identified Gaussian-shaped profile at time n=3. At time n=7 the variance of the profile is already close to that of the equilibrium  $F_\alpha$ , but its mean is still substantially shifted to the right. Thereafter, we observe the gradual motion of the profiles towards the left at a much lower speed. After 150 iterations, the approximate mean and variance of  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  are given by  $\approx 0.0474$  and  $\approx 1.8897$  respectively, where the latter agrees with the above exact value  $\sigma_\alpha^2$  up to 5 digits.

In Figure 7 we represent the convergence of the growth rates  $||F_n||_{L^1(\mathbb{R}^d)}/||F_{n-1}||_{L^1(\mathbb{R}^d)}$  towards the eigenvalue  $\lambda_{\alpha}$ . After 150 iterations we obtain that the growth rate takes the approximate value  $\approx 0.9857$ , which again agrees with the exact value  $\lambda_{\alpha}$  above up to 5 digits.

In Figure 8 we have represented the errors

$$\varepsilon_n^{\mathrm{mass}} := \left| \frac{\|F_n\|_{L^1(\mathbb{R}^d)}}{\|F_{n-1}\|_{L^1(\mathbb{R}^d)}} - \boldsymbol{\lambda}_{\alpha} \right|, \quad \varepsilon_n^{\mathrm{prof}} := \mathcal{D}_{\mathrm{KL}} \left( \frac{F_n}{\|F_n\|_{L^1(\mathbb{R}^d)}} \, \middle\| \, \boldsymbol{F}_{\alpha} \right),$$

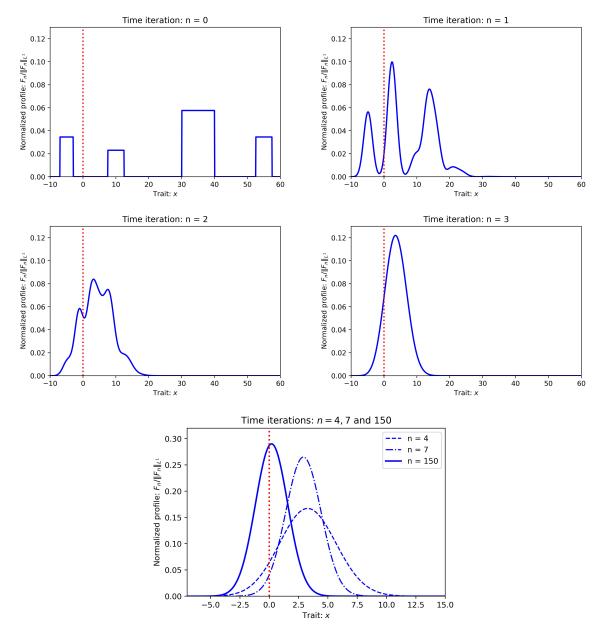


FIGURE 6. Relaxation of the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  towards the eigenfunction  $F_\alpha$  along the time iterations n=0,1,2,3,4,7 and 150 for a step function (7.1) as initial datum and weak selection parameter  $\alpha=0.015$ . The vertical dotted line represents the location of the mean of the equilibrium profile  $F_\alpha$ .

in semi-log plots so that the horizontal axis appears in the natural scale and represents each time iteration, and the vertical axis contains the logarithm of the errors. As observed, both plots reduce to essentially straight lines, suggesting exponential relaxation. Interestingly, the rate of convergence measured in Figure 8a is approximately 0.9442. It coincides numerically with the rate of relaxation among the subclass of Gaussian solutions (see Corollary 3.9), namely,  $\lambda_{\alpha}^4 \approx 0.9442$ , in contrast with the upper bound  $(2 \mathbf{k}_{\alpha})^{1/3} \approx 0.9905$  obtained in Theorem 1.1. These results are in perfect agreement with the other convergence results in Figure 8b: the numerical rate of convergence in the Kullback-Leibler divergence is 0.9441, which is close to the one among the class of Gaussian solutions  $\lambda_{\alpha}^4$ , and better than the upper bound obtained in Theorem 1.1, namely  $(2 \mathbf{k}_{\alpha})^{2/3} \approx 0.981$ . Moreover, the numerical rate of convergence of the variance of the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  is much faster: it is approximately  $\approx 0.4721 < 0.5$ , again close to the expected value for Gaussian solutions, namely,  $\mathbf{r}_{\alpha} \approx 0.472$ . These numerical results illustrate the two-step process arising in the relaxation dynamics: convergence towards a Gaussian profile occurs much

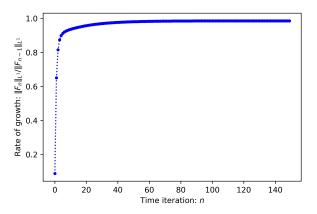


FIGURE 7. Relaxation of the growth rates  $||F_n||_{L^1(\mathbb{R}^d)}/||F_{n-1}||_{L^1(\mathbb{R}^d)}$  towards the eigenvalue  $\lambda_{\alpha}$  along time iterations  $0 \le n \le 150$  for a step function (7.1) as initial datum and weak selection parameter  $\alpha = 0.015$ .

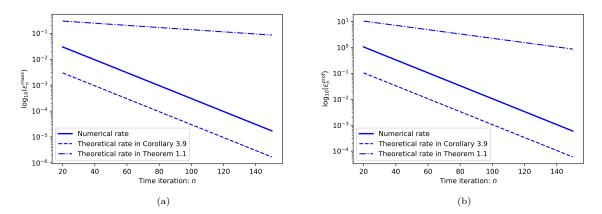


FIGURE 8. Numerical computation of the rates of convergence of the growth rates and the normalized profiles: (a) semi-log plot of the errors  $\varepsilon_n^{\text{mass}} := |\|F_n\|_{L^1}/\|F_{n-1}\|_{L^1} - \lambda_{\alpha}|$ , and (b) semi-log plot of the errors  $\varepsilon_n^{\text{prof}} := \mathcal{D}_{\text{KL}} \left( F_n / \|F_n\|_{L^1} \|F_{\alpha} \right)$ .

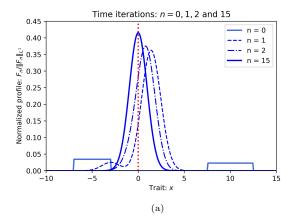
faster than relaxation of the mean towards the origin due to (weak) selection. Alternatively speaking, the equilibrium variance builds up much faster than the center of the distribution gets to the origin.

7.2. Strong selection. In this part, we consider a larger value of  $\alpha$ . We opted for  $\alpha = 0.4$ . Indeed, if  $\alpha$  is taken too large then there is no visual difference between the solution and the equilibrium configuration after one single iteration. This time, the numerical values of the features of the equilibrium are:

$$\lambda_{\alpha} \approx 0.7944, \quad \sigma_{\alpha}^2 \approx 0.9221.$$

The latter substantially differs from the value at linkage equilibrium (i.e.,  $\sigma_0^2 = 2$ ). In Figure 9a we note that the solution resembles a Gaussian distribution after a couple of iterations. After 15 iterations we obtain that the normalized profile  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  has approximate mean and variance respectively given by  $\approx 0.0017$  and  $\approx 0.9221$ . In Figure 9b we observe the convergence of the growth rates  $\|F_n\|_{L^1(\mathbb{R}^d)}/\|F_{n-1}\|_{L^1(\mathbb{R}^d)}$  towards the eigenvalue  $\lambda_{\alpha}$ . After 15 iterations the growth rate becomes  $\approx 0.7944$ , which again agrees with the exact value  $\lambda_{\alpha}$  above up to 5 digits.

Similar computations as in the previous Figure 8 allow finding numerical approximations for the rate of convergence of the growth rates and the normalized profiles. Specifically, we obtain an approximation  $\approx 0.3901$  for the rate of convergence of the growth rates. Once again, such a value is close to rate of relaxation among the subclass of Gaussian solutions, namely,  $\lambda_{\alpha}^4 \approx 0.3983$ , in contrast with the upper bound  $(2 \, \mathbf{k}_{\alpha})^{1/3} \approx 0.8578$  obtained in Theorem 1.1. Similarly, the numerical rate of convergence of the Kullback-Leibler divergence is 0.3966, which is close to the one among the class of Gaussian solutions  $\lambda_{\alpha}^4$ , and better than the upper bound obtained in Theorem 1.1, namely  $(2 \, \mathbf{k}_{\alpha})^{2/3} \approx 0.7358$ .



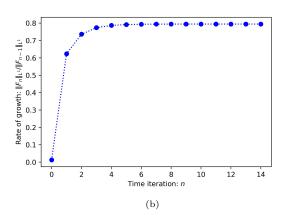


FIGURE 9. (a) Zoom near the origin of the relaxation of the normalized profiles  $F_n/\|F_n\|_{L^1(\mathbb{R}^d)}$  towards the eigenfunction  $\mathbf{F}_{\alpha}$  along the time iterations n=0,1,2 and 15 for the step function (7.1) as initial datum and strong selection parameter  $\alpha=0.4$ . The vertical dotted line represents the location of the mean of the equilibrium profile  $\mathbf{F}_{\alpha}$ . (b) Relaxation of the growth rates  $\|F_n\|_{L^1(\mathbb{R}^d)}/\|F_{n-1}\|_{L^1(\mathbb{R}^d)}$  towards the eigenvalue  $\lambda_{\alpha}$ .

#### 8. Conclusions and perspectives

In this paper, we have proven asynchronous exponential growth in a quantitative genetics model for the evolution of the distribution of traits in a population governed by sexual reproduction and multiplicative effect of selection. Our model assumes time-discrete non-overlapping generations, which rule out an eventual mixing with previous generations of ancestors. In addition, our non-linear sexual reproduction operator is set in agreement with Fisher's infinitesimal model, and we have chosen selection to act on the survival probability of individuals. Our main result provides quantitative convergence rates of the renormalized distributions towards a unique stationary profile. Indeed, rates are exponential, which can be interpreted loosely as a spectral gap in this non-linear context.

It is noticeable that the sexual reproduction operator is contractive under the Wasserstein distance (see Lemma 2.1). However, the generic incompatibility of multiplicative selection with transport distances becomes an apparent obstruction to the use of a direct perturbative approach in the regime of weak selection (see Example 2.4). This obstacle was circumvented by G. RAOUL [18], assuming that selection is restricted to a compact support. We follow a different route in a purely non-perturbative setting. To this end, we restrict to the specific choice of a quadratic selection function of the trait, for the sake of simplicity. Our alternative approach relies on an appropriate study of the propagation of information along large binary trees of ancestors, by a suitable reformulation of high-dimensional integrals, inspired by the changes of variables performed in [9] in the regime of small variance. This reveals an ergodicity property where the exact shape of the initial distribution is quickly forgotten across generations.

We remark that the above heuristic arguments indicate that the quadratic Wasserstein metric is not fully appropriate for dealing with the problem at hand. However, we could not identify yet a proper metric that extends the contraction property of the neutral case to the quadratic selection case.

Several perspectives are envisaged. First, there is an apparent price to pay with our method, in which we fully exploit the Gaussian structure induced by quadratic selection in order to perform tractable computations within the high-dimensional integrals. We believe though that the restriction to quadratic selection might be overcome in future works to allow for more general selection functions. Second, as studied in [9], the case of multiple minima on the selection function leads to non-uniqueness of stable equilibria. Then, uncovering the hidden metastability and quantifying the relaxation towards a specific equilibrium is of great interest for its applications in quantitative genetics. Finally, a major problem is to transcend non-overlapping generations and tackle the full time-continuous model as presented in previous literature.

#### Appendix A. Nondimensionalization and derivation of the time-discrete version

In this section, we shall nondimensionalize the time-continuous model (1.6)-(1.3) with Gaussian mixing kernel G and quadratic selection function m. By time discretization on the Duhamel formulation, we will derive the time-discrete version (1.1) which has been central in this paper. Indeed, we shall show that

we can reduce parameters into only one, namely, parameter  $\alpha \in \mathbb{R}_+$  in the quadratic selection function (1.5). Bearing in mind all the biological parameters and dimensions, our evolution problem reads:

$$\begin{cases}
\partial_t f = -r_m m(x) f + r_b \mathcal{B}_{\sigma^2}[f], & t \ge 0, x \in \mathbb{R}^d \\
f(0, x) = f_0(x), & x \in \mathbb{R}^d.
\end{cases}$$
(A.1)

Here,  $r_m, r_b \in \mathbb{R}_+$  represent the mortality and birth rates and have frequency units, whilst the dimensionless corrections of m and  $\mathcal{B}$  take the form

$$m(x) = \frac{|x|^2}{\sigma_m^2}, \qquad \mathcal{B}_{\sigma^2}[f](t,x) = \int_{\mathbb{R}^{2d}} G_{0,\sigma^2}\left(x - \frac{x_1 + x_2}{2}\right) \frac{f(t,x_1)f(t,x_2)}{\int_{\mathbb{R}^d} f(t,x') \, dx'} \, dx_1 \, dx_2,$$

for each  $x \in \mathbb{R}^d$ . Here,  $G_{0,\sigma^2}$  represents the Gaussian centered at 0 with variance  $\sigma^2$ , and  $\sigma_m$  and  $\sigma$  have the same units as the quantitative trait. On the one hand,  $\sigma_m$  can be regarded as a characteristic unit quantifying the effective range of selection. Namely, if  $|x| \geq \sigma_m$  then  $m(x) \geq 1$ . On the other hand,  $\sigma^2$  is the genetic variance and  $2\sigma^2$  represents the variance at linkage equilibrium often denoted by  $\sigma_{LE}^2 := 2\sigma^2$  (cf. [8]). We nondimensionalize our system by appropriately scaling our variables as follows

$$\widehat{t} = \frac{t}{T}, \quad \widehat{x} = \frac{x}{L}, \quad \widehat{f}(\widehat{t}, \widehat{x}) = L^d f(t, x).$$

Here, T and L are characteristic units for time and trait. For simplicity, we set them as follows in terms of the parameters of the system:

$$T := \frac{1}{r_m}, \quad L := \sigma.$$

In other words, we scale time according to the mortality rate  $r_m$  and the trait variable according to the genetic variance  $\sigma^2$ . Dropping hats for simplicity yields the following dimensionless form of the equation

$$\begin{cases}
\partial_t f = -\frac{\alpha}{2} |x|^2 f + \beta \mathcal{B}[f], & t \ge 0, x \in \mathbb{R}^d, \\
f(0, x) = f_0(x), & x \in \mathbb{R}^d,
\end{cases}$$
(A.2)

where  $\mathcal{B} = \mathcal{B}_1$  involves again unit genetic variance thanks to our scaling assumption. This reduces the amount of free parameters to only two of them  $\alpha, \beta \in \mathbb{R}_+$  defined as follows:

$$\alpha := \frac{\sigma_{LE}^2}{\sigma_m^2}, \quad \beta := \frac{r_b}{r_m}.$$

In the sequel, we justify our time-discrete version (1.11). Specifically, let us integrate equation  $(A.2)_1$  for  $t \in [t_n, t_{n+1}]$  where  $\{t_n\}_{n \in \mathbb{N}} \subseteq \mathbb{R}_+$  is any increasing sequence. Then, we recover Duhamel's formula

$$f(t_n, x) = e^{-\frac{\alpha}{2}|x|^2(t_n - t_{n-1})} f(t_{n-1}, x) + \beta \int_{t_{n-1}}^{t_n} e^{-\frac{\alpha}{2}|x|^2(t_n - s)} \mathcal{B}[f(s, \cdot)](x) \, ds,$$

for any  $n \in \mathbb{N}$  and each  $x \in \mathbb{R}^d$ . Using the rectangle rule for the integral in the right hand side leads to the approximations  $F_n(x) \simeq f(n,x)$  of the above time-continuous problem for f = f(t,x):

$$F_n(x) = e^{-\frac{\alpha}{2}|x|^2} F_{n-1}(x) + \beta e^{-\frac{\alpha}{2}|x|^2} \mathcal{B}[F_{n-1}](x), \tag{A.3}$$

for any  $n \in \mathbb{N}$  and each  $x \in \mathbb{R}^d$ . Here, we have set unitary time steps  $t_n - t_{n-1} = 1$  for simplicity and  $t_0 = 0$ . We emphasize that the right hand side of (A.3) consists of two different terms. On the one hand, the first term describes the amount of old individuals from the previously generation n-1 having resisted the effect of selection in generation n. On the other hand, the second term computes the amount of offspring conceived by individuals in the previous generation n-1, where an eventual decline due to selection has also been taken into account. Let us emphasize that in (A.3) individuals from generation n-1 can merge (after breeding) with their offspring to form next generation n. This is typical in most biological populations, which have overlapping generations (e.g., some multivoltine flies like Drosophila melanogaster). However, some other systems have non-overlapping generations so that the full generation n-1 gets extinct after breeding and only their offspring survive at generation n (e.g., some univoltine insects like the Dawson's burrowing bee). In that case, by canceling the first term (A.3) reduces to

$$F_n(x) = \beta e^{-\frac{\alpha}{2}|x|^2} \mathcal{B}[F_{n-1}](x),$$
 (A.4)

for any  $n \in \mathbb{N}$  and each  $x \in \mathbb{R}$ . We remark that two different parameters  $\alpha$  and  $\beta$  are still present in (A.4). However, given that  $F \mapsto e^{-\frac{\alpha}{2}|x|^2} \mathcal{B}[F]$  is a 1-homogeneous operator we can kill the parameter  $\beta$  by rescaling the trait distributions  $F_n$ . Specifically, define  $\widehat{F}_n = \beta^{-n} F_n$  and note that (A.4) reduces to our time-discrete problem (1.1) with only one parameter  $\alpha$ .

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