

# Data-driven modeling for river flood forecasting based on a piecewise linear ARX system identification

B. Hadid, E. Duviella, S. Lecoeuche

# ▶ To cite this version:

B. Hadid, E. Duviella, S. Lecoeuche. Data-driven modeling for river flood forecasting based on a piecewise linear ARX system identification. Journal of Process Control, 2020, 86, pp.44-56. 10.1016/j.jprocont.2019.12.007 . hal-03225015

# HAL Id: hal-03225015 https://hal.science/hal-03225015

Submitted on 21 Jul 2022

**HAL** is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers. L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.



Distributed under a Creative Commons Attribution - NonCommercial 4.0 International License

# Data-driven modeling for river flood forecasting based on a Piecewise Linear ARX system identification \*

Baya Hadid<sup>a,\*</sup>, Eric Duviella<sup>a</sup>, Stéphane Lecoeuche<sup>a</sup>

<sup>a</sup>IMT Lille Douai, University of Lille, Informatics and Automatics Research Unit, F-59000 Lille, France.

# Abstract

Most of the studies related to the rainfall-runoff modeling of rivers consist of data-driven models, given that the corresponding physical modeling approaches are based on a thorough geological knowledge of the river in addition to a time consuming simulation. Indeed, flood forecasting services have the difficult task of avoiding natural and human disasters and choose for that to use input-output or grey box models for their simplicity and easy calibration updates. However, these models are not evolving according to the variations of environmental conditions or need at least the evapotranspiration and the soil humidity measurements in addition to the rainfall quantity. This paper gives an alternative approach to the existing rainfall/runoff linear and nonlinear models by the utilization of a hybrid system consisting in a Piecewise Auto-Regressive eXogeneous (PWARX) structure identified using

Preprint submitted to Elsevier

December 26, 2020

<sup>\*</sup>This work has been supported by ARMINES through the contract n° 1908V/1700661and the IIW (Institution Intercommunale des Wateringues). Authors also thank the DREAL Hauts-De-France flood forecasting services (SPC) for providing access to data. Conflict of interest - none declared

<sup>\*</sup>Corresponding author

*Email addresses:* baya.hadid@imt-lille-douai.fr (Baya Hadid), eric.duviella@imt-lille-douai.fr (Eric Duviella), stephane.lecoeuche@imt-lille-douai.fr (Stéphane Lecoeuche)

an approach that alternates between data assignment and parameter estimation. The usage of this special kind of nonlinear systems bears a potential to handle the nonlinearities and varying-time delays mainly induced by the soil water storage and evapotranspiration.

*Keywords:* Rainfall-Runoff model, Hybrid system, Data-driven model, Non-supervised clustering, Data assignment

### 1 1. Introduction

Nonlinearities are the biggest challenge faced when one seeks for a com-2 plex large-scale natural system model that has to be reproducible, especially 3 when we have a limited knowledge of the intrinsic phenomena governing the 4 system and when this system is virtually unique, for example in terms of 5 geophysical properties. This is particularly true when it comes to model the rainfall-runoff relationship of a river for a precise flood forecast with having 7 as an only knowledge the precipitation past measurements and forecast and 8 the runoff measurements. Indeed, the generation process of a flood start-9 ing from precipitations in a natural catchment in order to deliver an early 10 warning for extreme hydrological events is one of the biggest challenges faced 11 by hydrologists. [1] assessed in its study at least 19 existing daily rainfall-12 runoff models, leading rapidly to a bigger number of models when it comes 13 to take into account the seasonality and the climate change we are currently 14 undergoing. The reason is that the stream level or discharge is not similarly 15 impacted by the same amount of rainfall according to seasons due to the 16 influence of the temperature on the evapotran piration of the ground and the 17 vegetative cover. Other factors affecting the runoff come into play such as 18

the soil permeability, the river slope, the drainage due to agriculture activity,
factors that differ from a river to another.

The common goal of all the rainfall-runoff modeling approaches is to pro-21 vide a flood forecast with a minimum lead-time (time between its announce-22 ment and arrival). The overwhelming majority of the proposed approaches 23 are either physical/mathematical approaches, data-based alternately named 24 black box approaches, and conceptual approaches. Mathematical/hydrological 25 modeling is known to be time consuming due to the huge number of param-26 eters to calibrate and it is thus critical to use for an online forecast. In 27 addition to the computational burden, it requires a non negligible amount of 28 detailed information on each concerned river such as a bathymetric survey. 20

On the contrary, data-driven, black box fully numerical modelling ap-30 proaches establish models by using only input and output measurements. 31 Many researchers have developed numerical runoff/rainfall models with vary-32 ing degrees of success. A good survey on non parametric data driven model-33 ing techniques is presented in [2]. Neural networks (ANNs), genetic program-34 ming, evolutionary polynomial regression, Support vector machines (SVM), 35 M5 model trees, K-nearest neighbors, and multiple linear regression tech-36 niques are implemented and evaluated using daily stream flow data. The 37 SVM approach was also explored in [3], where a short-term stream flow pre-38 diction was performed on hourly data. Other approaches based on the neural 39 networks were developed by for instance, introducing an *a priori* knowledge 40 on the evapotranspiration (see [4] and references therein). Neural Fuzzy 41 Networks (NFS) and preciseley ANFIS (Adaptive Network-based Fuzzy In-42 ference Systems) were also applied through the years ([5], wavelet neuro-43

<sup>44</sup> fuzzy models [6], and radial basis function artificial neural network extreme <sup>45</sup> training on monthly data in [7]) but still suffer from a non flexible structure <sup>46</sup> that depend on the expert's knowledge. Recent studies use adaptive fuzzy-<sup>47</sup> techniques for a more flexible structure achievement, such as the Self-adaptive <sup>48</sup> Fuzzy Inference Network (SaFIN) presented in [8]. However, and despite re-<sup>49</sup> spectable results provided by Machine learning techniques, they still need a <sup>50</sup> huge database to achieve a correct training and validation results.

In the family of black box parametric approaches, the linear models con-51 sisting in ARX and ARMAX models were firstly used due to their simplicity 52 [9, 10]. These methods have been abandoned because of the non linear-53 ities due to, as examples, evapotranspiration phenomenon and saturation 54 due to soil storage. Recently, a black box Linear Parameter Varying (LPV) 55 model was investigated for the Rainfall-Runoff Relationship (RRR) in urban 56 drainage networks [11] and rural catchment [12]. This kind of systems con-57 sider a lot of nonlinearities and depend on one or several external variables, 58 called scheduling variables, and then could be linearized at different operating 50 points resulting in a set of local Linear Time-Invariant (LTI) systems. The 60 issue with this approach comes from how to choose the right scheduling vari-61 able which is not trivial and could be a non observable variable or unknown 62 measurement. In [11], the scheduling variable was chosen as the output of a 63 non parametric model and the model was identified using the least-squares 64 algorithm when the scheduling parameter is taken as the output of the best 65 linear model in [12] but the optimal identification Simplified Refined Instru-66 ment Variable (SRIV) algorithm was applied. Both LPV methods lead to 67 acceptable results. In [13], the authors proposed an online recursive nonlinear 68

identification algorithm applied to Liane river (France) and was compared 69 to a recursive least-squares linear model over a future horizon of 24 hours. 70 Indeed, the online estimation allows to track the catchment intrinsic varia-71 tions by updating at each time-step the model parameters. The study over a 72 horizon is innovative comparing to the previous cited approaches and shows 73 that even the Fit score [14] or the Nash coefficient [15] are good, the intro-74 duction of a prediction horizon deteriorates the estimation results and then 75 increases the number of false alarms and missed alarms. 76

In [16], a conceptual daily lumped rainfall/runoff model called GR4J 77 (from the french "Génie Rural à 4 paramètres Journaliers") and its hourly 78 version GR4H, is presented as an improvement of the GR3J [17, 18] and the 79 performance was tested using five criteria. This conceptual model is based 80 on a reservoir system and its initialization requires the knowledge of the soil 81 saturation of the previous day. This approach has shown its reliability once 82 it was properly calibrated using a large amount of data and was applied on 83 several case studies [19, 20, 21]. 84

Hydromax tool used currently as a part of the Belgium forecasting sys-85 tem and presented in [22] represents the nonlinearities by a Hammerstein 86 structure using an *a priori* knowledge of the hydrological system to char-87 acterize the static function on the river basin maximum storage, the runoff 88 coefficient and the percolation parameter. An evapotranspiration estimation 89 is also required to calculate the effective rainfall which is the final input of a 90 predictive ARX model. Even if this approach is effective for important flow 91 forecasts, a larger prediction error is observed for low water. The authors 92 propose a combination or a switch between a linear ARX prediction model 93

#### <sup>94</sup> and nonlinear ARX (NARX).

In this paper, a rainfall-river level PieceWise Affine model (PWA) is pro-95 posed as a new solution to handle the maximum of nonlinearities. PWA sys-96 tems falls into the category of Hybrid systems. The term "hybrid" is related 97 to the interaction between continuous and discrete dynamics. The discrete 98 state sequence could be a result of the interaction between logic devices and 99 continuous processes or a physical phenomena that shows a discontinuous 100 behavior. A hybrid system could also be the result of an approximation of 101 a nonlinear system by a set of linear submodels. PWA model has the prop-102 erty of being universal due to its particular mapping and make this kind of 103 systems particularly attractive for nonlinear dynamical system identification 104 [23, 24].105

The principle is to consider the model as a combination of different sub-106 models. Each submodel is assimilated to a state or a mode and the switching 107 between the states depends for example on the change in the operating con-108 ditions. However, it is not obvious to know a priori the number of modes 100 and the state sequence, especially in case of natural large-scale systems. In-110 deed, it depends on the knowledge degree of our system and even so, there is 111 no guarantee to achieve satisfactory prediction performances. The literature 112 proposes multiple approaches of PWA estimation. These approaches can be 113 based on Algebraic procedures [25, 26], Bayesian procedures [27], Bounded-114 error (set-membership) procedures [28], Clustering procedures [29, 30, 31], 115 Optimization-based procedures [32, 33] Recursive procedures [34, 35, 36, 37]. 116 Since the objective is to model the rainfall-runoff relationship without a pri-117 ori knowledge on the physical system using real data provided by gauges, 118

it then becomes important to choose carefully the right method that makes 119 possible a PWA approximation with a minimum knowledge on the structure 120 especially on the number of modes. We prefer among them the unsupervised 121 techniques and precisely, the PWA model identification algorithm proposed 122 by [31]. This algorithm is an unsupervised clustering technique based on the 123 theory of evidence introduced by [38] and explored later by [39] for clustering 124 purposes on the basis of the evidence of nearby observation samples called 125 the k-nearest neighbors according to the euclidean distance between data. 126 It solves simultaneously the problem of data assignment to discrete modes, 127 the parameter estimation of submodel dynamics and the number of possible 128 regimes or modes, thanks to an evidential procedure. The unsupervised fea-129 ture of this approach allows therefore to do not make any assumption on the 130 total number of operating modes, only the linear model orders are fixed a131 priori. 132

In the algorithm scheme, each data is assigned and the number of sub-133 models is reduced at each iteration thanks to a criteria that combines the 134 prediction error and the Euclidean distance between data samples, and then 135 the parameters of each submodel are updated. The problem of the PWA 136 model estimation is thus solved by clustering data and estimating a parame-137 ter vector using data contained in each class and an identification technique of 138 a linear model such as the least squares algorithm, with ensuring a minimum 139 number of submodels that gives the better performances. This algorithm is 140 therefore an alternation between data assignment and parameter adaptation. 141 This paper is organized as follows: Section 2 introduces the PWA iden-142 tification problem statement with considering a short-term and a long-term 143

prediction horizon, states the Evidential algorithm principle and details the Dempster-Shafer rules used for data reassignment. The problem of mode sequence forecasting with a lead time considering validation data is addressed in Section 3 through the prism of the region estimation based on machine learning techniques. The water level forecasting results of a river located in the north of France are reported in Section 4. Finally, Section 5 summarizes the key findings and foreshadows possible improvements in future research.

### <sup>151</sup> 2. PWA with a prediction horizon identification problem statement

The basic principle of the proposed method is to put together data that are more likely to have been generated by the same and underlying linear submodel. For each group of data, we can hence estimate one parameter vector to represent the corresponding submodel. The method is based on a unsupervised clustering by alternating the regression data assignment to a submodel and the submodel estimation.

# 158 2.1. Problem description

Let us consider the following PieceWise AutoRegressive eXogenous (PWARX) model with a prediction horizon *H* 

$$y(k) = f_{\sigma_k}^{(k)}(\varphi_H(k)) + e(k), \qquad (1)$$

 $\sigma_k$  is the discrete state  $\sigma_k \in \{1, \ldots, s\}$  and  $\varphi_H(k)$  is the regression vector having the following structure

$$\varphi_H(k) = [y(k-H-1) \cdots y(k-H-n_a) u(k) u(k-1) \cdots u(k-H-n_b)]^{\top},$$
(2)

where  $u(k) \in \mathbb{R}^{n_u}$  and  $y(k) \in \mathbb{R}^{n_y}$  are respectively the input and the measured output of the system at time  $k \in \mathbb{Z}$ ,  $n_a$  and  $n_b$  are the orders of the discretetime ARX model transfer function and H is the prediction horizon. e(k) is a noise/error term regarded as a zero-mean Gaussian noise with a standard deviation  $\sigma^2$ . For the sake of exposition and according to the case study, only the Single Input Single Output (SISO) case is discussed in the following  $(n_u = n_y = 1)$ .  $\mathcal{F}$  is a piecewise linear map of the form

$$\mathcal{F}_{\sigma_k}^{(k)}(\varphi_H) = \begin{cases} \theta_1^\top \bar{\varphi}_H(k) & \text{if } \sigma_k = 1 \\ \vdots & \\ \theta_s^\top \bar{\varphi}_H(k) & \text{if } \sigma_k = s, \end{cases}$$
(3)

where  $\theta_{\sigma_k}$  is the parameter vector defining the linear submodel  $\mathcal{M}_{\sigma_k}$  and  $\bar{\varphi}_H(k) = \left[\varphi_H(k)^\top 1\right]^\top$  is the extended regression vector.

 $\sigma_k$  is described by a polyhedral partition of the regression space and is then defined by

$$\sigma_k = i \quad \text{iff} \quad \varphi_H(k) \in \mathcal{R}_i, \quad i = 1, \dots, s.$$
(4)

The regions  $\{\mathcal{R}_i\}_{i=1}^s$  form a complete polyhedral partition of the regression space  $\mathcal{R} \subset \mathbb{R}^n$  where  $n = n_a + n_b + 1$ . Each distinct region is described by

$$\mathcal{R}_i = \{\varphi_H(k) \in \mathbb{R}^n : \mathcal{H}_i \bar{\varphi} \le \mathbf{0}\}.$$
 (5)

 $\mathcal{H}_i \in \mathbb{R}^{(n+1)}$  is a function defining the separating boundaries delimiting the region  $\mathcal{R}_i$ . Finally, the piecewise affine map  $\mathcal{F}$  can be rewritten as follows

$$\mathcal{F}_{\sigma_{k}}^{(k)}(\varphi_{H}) = \begin{cases} \theta_{1}^{\top} \bar{\varphi}_{H}(k) & \text{if } \mathcal{H}_{1} \bar{\varphi}_{H} \leq \mathbf{0} \\ \vdots \\ \theta_{s}^{\top} \bar{\varphi}_{H}(k) & \text{if } \mathcal{H}_{s} \bar{\varphi}_{H} \leq \mathbf{0}. \end{cases}$$
(6)

The PWARX model identification consists in estimating the parameter vectors  $\{\theta_i\}_{i=1}^s$ , given a set of N observations  $\{\varphi_H(k)^{\top}, y(k)\}_{k=1}^N$  generated by the PWARX system defined by (1) and the boundaries' parameters of the partitions  $\{\mathcal{R}_i\}_{i=1}^s$  under the assumption that the discrete mode sequence  $\{\sigma_k\}_{k=1}^N$  and the number of submodels s are unknown.

PWARX model estimation is classically used with a null prediction horizon H = 0 and the algorithm is a one-step ahead prediction model. Since the objective is to forecast the output at a future time k + f at time k, noted k+f|k with  $f \ge 1$ , the prediction at time k+f with a null prediction horizon is given by

$$\hat{y}\left(k+f|k\right) = \hat{\theta}_{i}^{\top}\hat{\varphi}_{0}\left(k+f|k\right) \quad i = 1,\dots,s,$$
(7)

$$\hat{\varphi}_0 (k+f|k) = [ \hat{y}(k+f-1|k) \cdots \hat{y}(k+f-n_a|k) u(k+f) \cdots u(k+f-n_b) ]^\top, \quad (8)$$

where  $\hat{\varphi}_0$  is the regression vector containing the estimated output  $\hat{y}$  instead of the real output y. We note here that  $\hat{\varphi}_0$  contains the predicted outputs <sup>190</sup> previously and repeatedly estimated using the forecasted inputs u(k + 1)<sup>191</sup> to u(k + f - 1) to simulate the model output using  $\hat{y}(k) = \hat{\theta}_i^{\top} \hat{\varphi}_0(k)$ . The <sup>192</sup> repetition of this operation causes an error propagation. To achieve directly <sup>193</sup>  $\hat{y}(k + f|k)$ , a prediction horizon equal to H = f is used. The inconvenient <sup>194</sup> is that the model is more complex due to the higher number of parameters <sup>195</sup> to estimate. The predicted output is then defined by

$$\hat{y}\left(k+f|k\right) = \hat{\theta}_{i}^{\top}\varphi_{H}\left(k+f\right) \quad i = 1, \dots, s,$$
(9)

$$\varphi_H(k+f) = [y(k-1) \cdots y(k-n_a) u(k+f) \cdots u(k-n_b)]^{\top}.$$
 (10)

Increasing the prediction horizon leads then to an increasing of the re-196 gression vector dimension. However, the regression vector is commonly used 197 as feature vectors to estimate the regions boundaries based on a supervised 198 machine learning technique (see Section 3). It is yet recommended to keep 199 the dimension of feature vectors as low as possible to avoid simulation time 200 issues while ensuring that the clustering performances are not compromised 201 due to a poor dimension. On the other hand, increasing the prediction hori-202 zon limits the effects of the error propagation. A high training precision 203 remains however a necessity. This is possible thanks to a good value selec-204 tion of number of nearest neighbors which is the main tuning parameter of 205 the evidential algorithm defined in Section 2.2. 206

Based on this, we define first a new prediction horizon H' where  $1 \leq H' < H$  in Eq. (2) selected depending on the system dynamic and more

specifically smaller than the response time of the system. The predicted output is then simulated as for the case H = 0, *i.e.* applying repeatedly the operation  $\hat{y}(k + H') = \hat{\theta}_i^{\top} \hat{\varphi}_{H'}(k + H')$  with

$$\hat{\varphi}_{H'}(k+H') = [\hat{y}(k-1|k) \cdots \hat{y}(k-n_a|k) u(k+H') \cdots u(k+H'-n_b)]^{\top}.$$
(11)

Note that at each new prediction, the real values of the output are assimilated by replacing the estimated output by its real value in the regression vector  $\varphi_{H'}(k + H')$ .

#### 215 2.2. Algorithm principle and initialization

The approach is based on an unsupervised clustering combined with a linear regression to merge regression data belonging to the same submodel. In unsupervised clustering techniques, the class labels/indexes of regression data are not known *a priori*. They hence need a clustering rule or a similarity criterion to build the data partitioning.

The clustering based procedures consider that the local linearity characterizing the PWA map can be reconstructed using a resemblance between a submodel parameter vector and an estimated parameter vector generated from small local data sets. Each algorithm iteration consists of an alternation between regression data assignment to a class built using the unsupervised clustering and a submodel parameter estimation for this new updated class using identification approaches dedicated to linear ARX systems.

Since the number of submodels is unknown *a priori*, regression data are first partitioned into N clusters  $C = \{C_1, \ldots, C_N\}$ , where  $C_i = \{X(i)\}$ . We assign the initial parameter matrix  $\Theta^{(0)} = \left[\hat{\theta}_1^{(0)}, \ldots, \hat{\theta}_N^{(0)}\right]$  to these N clusters.

The objective is to reduce the number of clusters at each iteration and to 231 converge to s models by means of a decision rule that determine how data 232 migrate towards the most representative clusters. Thanks to the adaptation 233 procedure, the less representative clusters become empty and are eliminated 234 resulting in a decreasing number of clusters.  $\hat{\theta}_i$  is computed by considering 235 the data of the cluster  $C_i$  and its "c" nearest neighbors and by using a *least*-236 squares (LS) on these c + 1 data. Each data is achieved by concatenating 237 the regression vector and the observed output  $X(i) = \left[\varphi(i)^{\top}, y(i)\right]^{\top}, i =$ 238  $1,\ldots,N.$ 239

We briefly recall that a LS estimation based on an observed data sequence  $\{u(k), y(k)\}_{k=1}^{N_i}$  of the parameter vector defining the ARX system is given by

$$\hat{\theta}_i = \left(\sum_{k=1}^{N_i} \bar{\varphi}_H(k) \bar{\varphi}_H^\top(k)\right)^{-1} \sum_{k=1}^{N_i} \bar{\varphi}_H(k) y(k).$$
(12)

The tuning parameter c value is a compromise between the noise level 242 and the risk of achieving a number of mixed submodels generated from mixed 243 local data belonging to different submodels. If the noise level is high, its effect 244 is filtered by using a large value of c. In the case of rainfall/runoff observations 245 provided by gauges, the effect of noise is weak if not non-existent. A small c 246 is then recommended to avoid the problem of mixed submodels. In addition, 247 a small c value results in a higher number of submodels leading inevitably to 248 higher performances. Numerical examples show that beyond a certain value, 240 the number of submodels converges to the real s beyond a certain value [31]. 250

### 251 2.3. Data reassignement based on Dempster-Shafer theory

Dempster-Shafer theory, also called theory of Belief functions, is used to 252 model information uncertainty [38]. This theory is thus based on the belief 253 functions that allows reasoning on uncertain facts, not relying on probabilistic 254 quantification, but in a more general way than the Bayesian model. The 255 Transferable Belief Model (TBM) represents the quantified beliefs of an agent 256 and was subject to varying interpretations [40]. The Dempster-Shafer theory 257 is composed of a first level called *credal* where the beliefs are quantified and 258 merged and a second level called *piqnistic* which comes from the Latin *piqnus* 259 which means literally *a bet* where decisions are made by a transformation of 260 the belief functions to probability functions. 261

Let  $\mathcal{C}$  a set of propositions called a frame of discernment and A a subset of  $\mathcal{C}, A \subseteq \mathcal{C}, m(A)$  is a part of the belief in the membership of the agent Xto the subset  $A. m: \mathcal{C} \to [0, 1]$  is defined as follows

$$m\left(\emptyset\right) = 0\,,\tag{13a}$$

$$\sum_{A \subseteq \mathcal{C}} m\left(A\right) = 1. \tag{13b}$$

The m(A) values are called the *Basic Belief Masses* (BBM) and the mfunction is the *Basic Belief Assignment* (BBA). As an example, a total lack of information concerning the potential membership of the agent X to a subset of  $\mathcal{C}$  is represented by m(A) = 0 for any strict subset A of  $\mathcal{C}$ .

If we transfer this interpretation to the case of the nearest neighbors for the data assignment, we define  $\Gamma_c(X(i))$  as a set of the *c* nearest neighbors

of X(i). Each neighbor  $X(j) \in \Gamma_c(X(i)), j = 1, \ldots, c$  can be considered as 271 a piece of evidence that influences one's belief in the cluster  $\mathcal{C}$  membership 272 of the data X(i), and a mass of belief  $m_j^i$  is associated to their relationship. 273 In other words, if X(j) is a member of the class  $C_P$ ,  $P \in \{1, \ldots, \bar{s}\}$  where 274  $\bar{s}$  is the estimated number of submodels, then a part of the mass of belief 275 is allocated to the action of assigning the data point X(i) to the cluster  $\mathcal{C}_P$ 276 given by  $m_i^i(\{\mathcal{C}_P\})$  and the remaining part, to the set  $\mathcal{C}$  as described as 277 follows 278

$$m_{j}^{i}(A) = \begin{cases} \alpha_{0}\phi_{j}^{i} & \text{if } A = \mathcal{C}_{P} \\ 1 - \alpha_{0}\phi_{j}^{i} & \text{if } A = \mathcal{C} \\ 0 & \text{if } A \in 2^{\mathcal{C}} \setminus \{\{\mathcal{C}_{P}\}, \mathcal{C}\} \end{cases}$$
(14)

279 where

$$\phi_j^i = \exp\left(-\gamma_P \|X(i) - X(j)\|^2 - \beta_P \left(y(i) - \theta_P^\top \bar{\varphi}_H(i)\right)^2\right).$$
(15)

What is notable in  $\phi_j^i$  expression, is that the exponent is composed of two terms. The first one expresses the Euclidean distance between X(i) and the neighbor X(j) represented by  $d_j^i = \sqrt{(X(i) - X(j))^\top (X(i) - X(j))}$ , and the second term represents the prediction error related to the estimation of the parameter vector  $\theta_P$  of the submodel P, associated to the cluster  $\mathcal{C}_P$ . This formulation aims at minimizing simultaneously the euclidean distance and the output error (both shown in Figure 1.

 $\alpha_0$  is a positive parameter such as  $0 \ll \alpha_0 \ll 1$  that prevents from allocating the entire mass of belief to the class  $C_P$ . Indeed, even if there is zero distance between X(i) and X(j) and the prediction error is also towards zero, it remains an uncertainty concerning the belonging to the same class.



Figure 1: Example of three affine submodels in the augmented regression space

The mass of belief of X(j) on cluster  $C_p$  *i.e.*  $m_j^i(\{C_P\})$  is therefore function of the mean distance  $d_P$  between two data belonging to the same class  $C_P$  and its proximity with the linear model expressed by the average error  $e_P$  between the measured output and the submodel output as described in the following

$$\beta_P = \frac{1}{d_P^2},\tag{16a}$$

$$\gamma_P = \frac{1}{e_P^2}.\tag{16b}$$

Simple heuristics for the choice of  $\alpha_0$  and the initial value of  $\gamma_P$  is presented in Section 4.2. One can note that if X(i) is far from X(j) and in the same time the error  $(y(i) - \theta_P^\top \bar{\varphi}(i))^2$  is large, the class  $\mathcal{C}_P$  of X(j) will be considered as providing very little information regarding the class of X(i).

For more detailed and comprehensive explanation on how the decision rules are built based on the combination between the belief functions and <sup>297</sup> Dempster's rule, the reader is invited to refer to [39, 31, 41].

# 298 Credal level:

After tuning the c BBAs  $m_1^i, m_2^i, \ldots, m_c^i$ , the combination of all the BBAs using Dempster's combination rule is performed using the orthogonal sum  $\oplus$ and yields to a single mass

$$m^i = m_1^i \oplus \ldots \oplus m_c^i, \tag{17}$$

which can be calculated using a second rule called the conjunctive rule of combination. In case of A = C, it is defined by

$$m^{i}(\mathcal{C}) = \sum_{A_{1}, A_{2}, \dots, A_{c}| \bigcap_{j} A_{j} = \mathcal{C}} \prod_{j=1}^{c} m_{i}^{j}(A_{j}).$$
(18)

It is obvious that:  $\left(\bigcap_{j} A_{j} = \mathcal{C}\right) \Rightarrow A_{j} = \mathcal{C}$ . According to Eq. (14),  $m^{i}(\mathcal{C})$ becomes

$$m^{i}(\mathcal{C}) = \prod_{j=1}^{c} (1 - \alpha_{0} \phi_{j}^{i}).$$
(19)

The same reasoning applies to the belief function  $m^i(\{\mathcal{C}_P\})$  which is given by

$$m^{i}(\{\mathcal{C}_{P}\}) = (1 - \prod_{j/X(j)\in\mathcal{C}_{P}} (1 - \alpha_{0}\phi_{i}^{j})) \prod_{\substack{q=1\\q\neq P}}^{\bar{s}} \prod_{j/X(j)\in\mathcal{C}_{P}} (1 - \alpha_{0}\phi_{i}^{j}).$$
(20)

<sup>308</sup> Pignistic level:

Let  $\mathcal{P}$  a set of probability distributions derived from the BBAs. the unknown pignistic probability measure  $\mathcal{P}$  satisfies

$$Bel(A) < \mathcal{P}(A) < Pl(A),$$
 (21)

where Bel(A) is the belief function also called the credibility of A and Pl(A)is the plausibility of A defined respectively by

$$Bel^{i}\left(\{\mathcal{C}_{P}\}\right) = m^{i}\left(\{\mathcal{C}_{P}\}\right), \qquad (22a)$$

$$Pl^{i}\left(\{\mathcal{C}_{P}\}\right) = m^{i}\left(\{\mathcal{C}_{P}\}\right) + m^{i}\left(\mathcal{C}\right).$$
(22b)

313  $\mathcal{P}^i$  is then expressed by

$$\mathcal{P}^{i}\left(\{\mathcal{C}_{P}\}\right) = m^{i}\left(\{\mathcal{C}_{P}\}\right) + \frac{m^{i}(\mathcal{C})}{\bar{s}}.$$
(23)

The decision is made by assigning the data X(i) to the cluster  $C_q$  with maximum of pignistic probability. Then the decision rule is given by:

$$C_q = C_q \cup \{X(i)\}$$
 such that  $q = \max_{P,P=1,\dots,\bar{s}} \left(\mathcal{P}^i\left(\{\mathcal{C}_P\}\right)\right)$ . (24)

An iteration procedure allows the convergence of the clusters and their parameters using a stop criterion based on a comparison between the old and the new parameter vectors. Indeed, an adaptation of the parameter vector  $\hat{\theta}_P$  is performed at each new iteration *it* after a new data reassignment. The clustering is stabilized if

$$\left\|\Theta^{it+1} - \Theta^{it}\right\| \le \epsilon,\tag{25}$$

<sup>321</sup> where  $\epsilon$  is fixed by the user.

Finally, all the previous steps are summarized in Algorithm 1.  $\beta_0$  is ini-322 tialized using Eq. (16a),  $\gamma_0 = \nu \beta_0$  where  $\nu$  is a positive coefficient arbitrarily 323 chosen by the user. This parameter allows to adjust the weights in each of 324 the two terms of the exponent in Eq. (15) corresponding to the euclidean 325 distance and the submodel error to give each more or less importance de-326 pending on the case study. The parameters  $\gamma_P$  and  $\beta_P$  are adapted after 327 each iteration in order to take into account the evolution of the clusters. The 328 number of the nearest neighbors c is the main tuning parameter. After a 329 finite number of iterations, the procedure converges relatively fast to a rea-330 sonable number of submodels.  $\bar{s}$  is the number of non empty clusters but, in 331 the case of noisy data, the cardinality of some clusters falling under a given 332 number are removed. The problem of noisy data does not generally arise 333 in case of measurements provided by hydrometric stations. However, some 334 outliers could be encountered due to defective or blocked sensors. 335

#### **336 3.** Model validation for event forecasts

The computation of the river flow forecasts over a prediction horizon H involves necessarily the knowledge of the discrete state with a time-step ahead equal to this prediction horizon. In fact, the issue of the estimation of the discrete state on validation data was rarely addressed in literature. The literature survey is even rarer when it comes to real case studies, especially

# Algorithm 1

# 1: Initialization

- Set c,  $\alpha_0$ ,  $\beta_0$ ,  $\gamma_0$ ,  $\bar{s} = N$ , it = 0 and  $\epsilon$  (e.g.  $\epsilon = 10^{-4}$ ).
- Set  $C_i = \{X(i)\}, i = 1, ..., \bar{s}.$
- Estimate  $\Theta^{(0)} = \left[\hat{\theta}_1^{(0)}, \dots, \hat{\theta}_{\bar{s}}^{(0)}\right].$

# 2: Data reassignment

for i = 1, .., N

- For all  $X(j) \in \Gamma_c(X(i)), j = 1, \dots, c$ , calculate  $\phi_i^j$ , Eq. (15).

- Calculate  $m_j^i(\{\mathcal{C}_P\})$  and  $m_j^i(\{\mathcal{C}\})$ , Eq. (14).
- Combine all the BBA functions using Dempster's rule, Eq. (19) and (20).
- Calculate all the  $\mathcal{P}^i(\{\mathcal{C}_P\})$  using Eq. (23).
- Decide on the assignment of X(i) based on Eq. (24).

# end for.

- $\bar{s}$  = number of non empty clusters.
- Adaptation of the parameters:
  - Adapt  $\Theta^{(it)}$  using LS algorithm on the data of each non empty cluster.
  - Adapt the parameters  $\beta_P$  and  $\gamma_P$ , Eq. (16a) and (16b).

# 3: Convergence test

```
If \left\| \Theta^{(it+1)} - \Theta^{(it)} \right\| \leq \epsilon

s = \bar{s}.

else

Set it = it + 1 and return to step 2.

end.

4: Return \Theta^{it} and \hat{\sigma}(k).
```

when the number of states is relatively important, thus making the task of
state prediction even more delicate. Also, the case of discrete state prediction

over a prediction horizon was not investigated to the best of the author'sknowledge.

The output of the data classification step provides the discrete state es-346 timation  $\hat{\sigma}(k)$  and the estimates of the submodel parameters  $\Theta$ . It is now 347 possible to estimate the shapes of the polyhedral region boundaries. These 348 regions are in reality hyperplanes that form a complete polyhedral parti-349 tion  $\{\mathcal{R}_i\}_{i=1}^s$  of the regression space. The problem can then be equivalent 350 to the separation to s sets using linear classifiers [24]. However, the same 351 author highlights two problems encountered when a linear classification is 352 used. First, it is not possible to exactly estimate the regions starting from a 353 finite set of data which induces small errors in shaping, leading to the second 354 problem consisting in a misclassification of the regression vector located near 355 the discontinuities and larger prediction error can then be observed. A pos-356 terior re-attribution to the correct submodel could be envisaged during the 357 validation. For a more complete discussion about class separation methods 358 applied for SARX and PWARX systems, the reader is referred to Section 4.2 350 of [24]. 360

The Multicategory Support Vector Machine (SVM) algorithm [42, 43] is used to predict the shapes of partitions. This technique is commonly used in literature related to PWARX identification approaches in order to estimate the polyhedral regions. In [44], a SVM classifier using a one-versus-the-rest approach was applied.

# <sup>366</sup> 4. Application to flood forecasting

# 367 4.1. Case study

Our study focuses on the Liane coastal river of  $37 \ km$  long. It is situated in 368 the north of France (see Fig. 2) and flows in the La Manche sea (the English 369 Channel as called by the British). The catchment drained by this river covers 370 a total area of 244  $km^2$  and flows out a significant annual amount of rainfall 371 higher than 480 mm, which is much greater than the national average. The 372 Liane river has an average annual flow of 1.81  $m^3/s$  [45] corresponding to 373 a stream level of 47 mm but knows an increasingly number of severe flood 374 events due to the drainage network practiced in agriculture and a limited 375 exchange between the ground and surface water, exceeding frequently the 20 376  $m^3/s$ . We can cite for example the flood of November 2012 where its level 377 achieved 4,37 m. 378

A hydrometric station is situated in Wirwignes (red marker in Figure 2) 379 and is equipped with a telemetered rain gauge since 2003 and a limnimeter 380 since 1970. The rain gauge gives an information on the quantity of rain falling 381 on the catchment and the limnimeter gives the measurement of the river 382 upstream level. The data are collected with a basic time-step of 1 hour. It is 383 important to note that most of the studies related to flood forecasting deal 384 with the prediction of the streamflow instead of the water level. However, 385 in practice, the streamflow is not directly measured but is inferred using 386 a periodically revised relation between the measured stream level and the 387 streamflow. This relation is determined by a simultaneous measurements of 388 these two physical quantities over the natural range of flows from the lowest 389 to the highest values corresponding to the floods. 390



Figure 2: Localisation of Liane river and the hydrometric station (red marker) in the north of France.

Figures 1.a and 1.b shows respectively the hourly rainfall and the hourly 391 water level measurements, from January 2010 to June 2018 i.e. over a period 392 of eight years and a half. Two high water level periods are noticed: during 393 winter and at beginning of spring. The low level period is during summer 394 and autumn. The Liane is also characterized by a short hydrological response 395 time and a flood period less than 24 hours. The objective of the study is to 396 forecast the river level with the prediction horizon of at least 24 hours, using 397 the rainfall forecast provided by weather services and the observed stream 398 levels. 399



Figure 3: Model input and output from January 2010 to June 2018. (a) Rainfall in [mm] (b) Liane level [m]. The blue dashed line refers to the green threshold.

# 400 4.2. Prediction results

The Liane river underwent from January  $1^{st}$ , 2010 to June  $13^{th}$ , 2018, 55 401 significant floods where the river level exceeds the threshold of 2 m which 402 corresponds to a return period of one year (the green threshold, see Fig. 403 3) which still relatively a low threshold considering that in practice, it is 404 taken equivalent to a return period of two years (yellow threshold). Since we 405 are interested in the flood prediction, only the flood periods corresponding 406 to heavy rainfall seasons are manually selected for the study. Each period 407 duration is between 3000 and 5000 hours. Hence, only 31072 samples on 408 74045 are selected. About 70% of the 55 most significant events are used 409 for the training (which corresponds to 40 events and 20715 samples) and 410 the remaining events are used for validation (15 events and 10357 samples). 411 Figure 4 represents level samples used for estimation (black line) and for 412

<sup>413</sup> validation (red line).



Figure 4: Training data (black line) together with checking data (red line)

Different criteria are used to globally appreciate the performances of each model. The *FIT* criterion introduced by Ljung [14] defined by (26), the Mean Square Error defined by (27) and the Nash-Sutcliff coefficient [15] defined by (28) give an information about the global fitting of the predicted output in the observed data, considering that a prediction is performed each f points (of the f next water level samples).

$$FIT = 100 \times (1 - \frac{\|y - \hat{y}\|_2}{\|y - \overline{y}\|_2}),$$
(26)

$$MSE = \frac{1}{N} \sum_{k=1}^{N} (y(k) - \hat{y}(k))^2, \qquad (27)$$

$$NASH = 1 - \frac{\frac{1}{N} \sum_{k=1}^{N} \left[ \hat{y}(k) - y(k) \right]^2}{\operatorname{var}(y(k))},$$
(28)

where  $\overline{y}$  is the mean of the observed output, N is the total number of samples and var(.) is the signal variance. Other indicators are used specifically to the flood events in hydrological modeling. The most important one is the *Critical Success Index CSI* [46] which represents the ratio between the number of "Correct alarms" *CA* which are the number of the well-estimated crests with respect to the total number of alarms *i.e.* the significant events. *CSI* is defined by (29).

$$CSI = \frac{CA}{CA + MA + FA} \times 100.$$
<sup>(29)</sup>

If the observed event is not simulated or underestimated with more than 10% of the crest maximum, a Missed Alarm occurs and MA is the number of Missed Alarms. Conversely, a False Alarm occurs if a non observed event is simulated or overestimated with more than 10% and FA is the number of False Alarms. The MA and FA indicators are then defined as follows for each significant event

433 If  $|y(k) - \hat{y}(k|k - f)| > 0.1 \times |y(k)|$ :

$$\begin{cases} MA: & \text{if } y(k) > \hat{y}(k|k-f) \\ FA: & \text{if } y(k) < \hat{y}(k|k-f) \end{cases} .$$

$$(30)$$

We thus simply define a CA for events fulfilling the following condition :  $|y(k) - \hat{y}(k|k - f)| \le 0.1 \times |y(k)|.$ 

#### 437 4.2.1. Short-term forecasting

The selection of the short-term prediction horizon depends usually on the 438 sample time and the response time of a unit hydrograph. A typical choice 439 of the natural response time is one fifth to one third of a unit hydrograph 440 peak time in order to consider the slightest dynamic changes [22]. However, 441 since SVM technique is used to predict the discrete sequence mode, it is 442 recommended to shorten the prediction horizon to ensure a better validation. 443 Indeed, the dimension of the features used for training which consists in the 444 regression vectors depends on the horizon value. Higher is H, higher is 445 the dimension of features and more restrictive is the data recognition. A 44F classical solution commonly used in machine learning algorithms consists in 447 reducing the features dimension by replacing the regression vector by some 448 statistical metrics such as the mean, the variance and the kurtosis values. 449 The problem of this procedure is the risk of loss of information related to the 450 exact location of the local data provided initially by the regression vectors 451 leading to a confusion between submodels resulting in different data located 452 in the same region with same parameter vectors. 453

The peak time of the Liane river is about twelve hours which implies a response time between two and four hours. Finally, the prediction horizon is reduced to one hour for the previously stated reasons.

Model orders for the tested methods were chosen according to the performances structures built by increasing  $n_a$  and  $n_b$  from 1 to 20. Due to a high signal to noise ratio, it is allowed to choose the tuning parameter c small relatively to the number of training data samples to avoid a mixed submodel stemming from an important number of mixed local data as mentioned in 462 the previous section.

The choice of the parameter c is difficult to make *a priori*, especially when the method is sensitive to this parameter. A systematic search is necessary to obtain optimal results with a minimum number of modes.

To see the influence of the parameter c on the estimation of the number of submodels and for a selection that will approach the optimum as near as possible, the PWARX identification algorithm is applied to the same training set for different c values. The *FIT* score is calculated after the evaluation of the PWARX model on the validation set and for a long-term forecasting chosen arbitrarily equal to f = 24.

### 472 Selection of the hyperparameters:

Fig. 5 shows the resulting number of modes and FIT scores on validation data for f = 24 with c ranging from 50 to 500. It can be noticed that, non surprisingly, the smaller the c value, the larger is the number of estimated modes and FIT score. However, FIT scores for c = 150 and c = 200 are substantially equal to each other whereas the resulting number of modes are respectively  $\bar{s} = 85$  and  $\bar{s} = 60$ . The choice was naturally directed towards c = 200.



Figure 5: Number of submodels  $\bar{s}$  estimated by the Evidential algorithm and the *FIT* resulting values on validation data set for f = 24 as a function of c nearest neighbors.

 $\gamma$  is usually chosen between 0.1 and 5 and  $\beta$  is taken 10 to 40 times 480 greater then  $\gamma$ . The ratio between these two weighting parameters is in 481 fact rather more important then the values themselves given that the goal 482 is to favour the prediction error in this case study rather than the euclidean 483 distance in Eq. (15). These hyperparameter c gives the user the possibility 484 to achieve a high accuracy depending on the application but at the expense 485 of an important number of resulting modes. It is although possible to achieve 486 equivalent accuracy values with a smaller number of modes by testing one of 487 the parameters  $\beta$  or  $\gamma$ . 488

The best PWA structure in terms of performances criteria using estimation data is for  $n_a = 1$ ,  $n_b = 1$ , c = 200,  $\gamma = 0.5$ ,  $\beta = 20$ , leading to the training scores FIT = 93.7%, Nash = 0.996 and 60 classes. Fig. 6 shows the simulated output, the prediction error and the estimated discrete state

### <sup>493</sup> resulting from the training step.



Figure 6: (a) Observed level used to train the PWARX model (black solid line) together with PWA model output (red solid line) (b) Prediction error (c) Predicted discrete state, for H = 1 and f = 1

The Evidential algorithm gives the number of modes and the discrete 494 state sequence. The latter information is considered as a labelling of data 495 and is used to train a Support Vector Machine (SVM) classifier [42] in order 496 to predict the discrete state sequence corresponding to the validation data. 497 Although the fact that the model is piecewise linear, the kernel function 498 used for training is the RBF (Radial Basis Function) instead of a linear 499 function to minimize a wrong classification of the regression vectors close 500 to linear boundaries (see Section 3). The normalized regression vectors are 501 used as inputs of the SVM classifier. The feature vectors are normalized 502 by subtracting the mean in the numerator and dividing by the standard 503

deviation in order to make the values of each feature have zero-mean and unit-variance [47]. Feature normalization is a classical preprocessing step in machine learning that is required when features have different ranges and it has shown in our case study better performance results.

In order to assess the effectiveness of the classification, some measures 508 based on the Confusion Matrix are calculated: the Accuracy rate, the Recall 509 rate and the Precision rate. We briefly recall that a Confusion Matrix [48, 49] 510 is a concept coming from machine learning. It contains information about 511 actual and predicted classifications and is then used to estimate the overall 512 classification accuracy. A confusion matrix is a two-dimensional matrix, one 513 is indexed by the actual class of an object, the other is indexed by the class 514 that the classifier predicts. 515

Table 1 presents the basic form of confusion matrix for a multi-class classification task, with the classes  $C_1, \ldots, C_{\bar{s}}$ .  $N_{ij}$  represents the number of misclassed data *i.e.* belonging to class  $C_i$  but classified as class  $C_j$ . The confusion matrix is computed using the matlab routine "confusionmat".

		Predicted						
		$\mathcal{C}_1$	• • •	$\mathcal{C}_j$	•••	$\mathcal{C}_{ar{s}}$		
	$\mathcal{C}_1$	$N_{11}$	•••	$N_{1j}$	•••	$N_{1\bar{s}}$		
ual	:	:		÷		÷		
Act	$\mathcal{C}_i$	$N_{11}$		$N_{ij}$	• • •	$N_{i\bar{s}}$		
	:	÷		÷		÷		
	$\mathcal{C}_{ar{s}}$	$N_{\bar{s}1}$		$N_{\bar{s}j}$	•••	$N_{\bar{s}\bar{s}}$		

Table 1: Confusion matrix

Accuracy is the proportion of the total number of the correct predictions and is then defined as the ratio between the trace and the sum of the confusion 522 matrix as follows

$$Accuracy = \frac{\sum_{i=1}^{\bar{s}} N_{ii}}{\sum_{i=1}^{\bar{s}} \sum_{j=1}^{\bar{s}} N_{ij}}.$$
 (31)

Recall is a measure of the ability of a prediction model to select instances of a certain class from a data set. The overall Recall rate is the average value of each individual recall rate and is provided by the following formula

$$Recall = \frac{1}{\bar{s}} \sum_{i=1}^{\bar{s}} \frac{N_{ii}}{\sum_{k=1}^{\bar{s}} N_{ik}}.$$
(32)

Precision is a measure of the accuracy provided that a specific class has been predicted. The overall Precision rate is the average value of each individual class precision rate and is provided by

$$Precision = \frac{1}{\bar{s}} \sum_{i=1}^{\bar{s}} \frac{N_{ii}}{\sum_{k=1}^{\bar{s}} N_{ki}}.$$
(33)

The resulting SVM Accuracy indicator is 91.78%, the Recall indicator is 88.48% and the Precision indicator is 88.81%, showing a good recovery of the estimated state sequence.

Table 2 compares the PWA performance indicators with one-step ahead prediction horizon (H = 1, f = 1) with the LPV approach on the checking data *i.e.* the 15 remaining flood events. The LPV approach is chosen for comparison to PWARX modeling instead of the Nonlinear ARX (NARX) which is the other data-driven nonlinear approach because of the difficulty encountered to opt for the most effective static nonlinearities of both inputs and outputs. Several static functions were tested (linear, Radial Basis
Function with two tuning parameters, logarithmic, sigmoid, etc) and none of
them was able to give fair results. Bad performances would necessarily be
attributed to a bad selection of the static functions.

The index in  $LPV_1$  and  $PWA_1$  is relative to f = 1. LPV model struc-542 ture is chosen according to the performance indicators for  $n_a = 1, \ldots, 10$ 543 and  $n_b = 1, \ldots, 10$ . Concerning the LPV model, four different possibilities 544 for the scheduling variable are tested: past rainfall samples (u(k-1), u(k-1), u(k-1545  $(2), \ldots, u(k-24))$ , past water level samples  $(y(k-1), \ldots, y(k-5))$ , linear 546 model output [50], identified using for example Least Squares algorithm or 547 an Output Error (OE) algorithm based on a nonlinear minimization of the 548 prediction error based on Levenberg-Marquardt algorithm [51, 52], or the Re-549 fined Instrumental Variable identification of a Box-Jenkins Transfer Function 550 model [53]. The last possibility is a combination between the average of past 551 rainfall inputs and past water level  $([mean(u(k-1), \ldots, u(k-i)) y(k-1)])$ 552 where i = 1, ..., 24). A polynomial dependency function to the scheduling 553 variable with a polynomial degree  $r = 1, \ldots, 5$  is tested [10]. The OE ap-554 proach based on Levenberg-Marquardt algorithm is extended to the discrete-555 time LPV case and used for LPV model identification. PWA model outputs 556 are presented in Fig. 7. 557

	% FIT	Nash	MSE	Max	CA	MA	FA	% CSI
$PWA_1$	90.9	0.9917	$9.6 \times 10^{-4}$	0.99	15	0	0	100
$LPV_1$	94.7	0.997	$3.3 \times 10^{-4}$	0.47	15	0	1	93.75

Table 2: Performance results of the tested models on validation set, for H = 1, f = 1



Figure 7: (a) Actual level (black solid line) together with PWA model output (red solid line) (b) Prediction error (c) Predicted discrete state, with H = 1 and f = 1

According to Table 2, the performance indicators are in favor of the LPV 558 model for a very short-term forecast f = 1, albeit with one FA. The good 559 performances are most observed for lower water levels. One can note that 560 the PWA performance indicators did not significantly decrease in validation. 561 This is primarily due to a good selection of training data and tuning param-562 eters, and obviously, to the small prediction horizon value, thus limiting the 563 loss of accuracy in the SVM prediction of the discrete state and the error 564 accumulation due to the multiple iterations. 565

# 566 4.2.2. Long-term forecasting :

<sup>567</sup> In case of prediction horizons larger than the natural response time, a <sup>568</sup> good forecast of the precipitation is required and must be provided by the

user using weather forecasts services. In the following, an assumption on the 569 exact knowledge of the future rainfall is made. Naturally, a discussion on a 570 pessimistic and optimistic forecast must be conducted to quantify to what 571 extent the model is sensitive to its inputs. A comparison between the LPV 572 model and the PWARX model for 6h, 12h, 24h and 48h prediction horizons 573 is driven and the performance indicators are drawn in Table 3 and it shows 574 the clear superiority of the PWA performance indicators regardless to the 575 lead time. 576

In addition, the percentage of forecasted warnings regarding the green 577 threshold overrun with a temporal prediction accuracy smaller than 1 hour 578 is 88% for a prediction horizon H = 24, making this approach one of the best 579 ARX approaches in terms of prediction lag performances. The peaks are also 580 globally forecasted in time where the peak prediction lag percentage smaller 581 then 1 hour is 85%. Note that the CSI is also a good indicator of the peak 582 prediction temporal accuracy since it is impacted in case of an important 583 lag between real and forecasted peak. Fig. 8 shows the water level and the 584 prediction error when considering 24h ahead PWA forecasts. 585

		$\% \ FIT$	Nash	MSE	Max	CA	MA	FA	$\% \ CSI$
LPV	H = 6	67.5	0.894	0.0123	1.74	10	5	1	62.5
	H = 12	57.03	0.815	0.021	1.96	3	11	1	20
	H = 24	50	0.75	0.029	1.88	2	12	1	13.3
	H = 48	46.4	0.713	0.033	1.64	1	14	0	6.7
PWA	H = 6	85.5	0.9790	0.0025	1.42	15	0	1	93.75
	H = 12	84.7	0.9767	0.0027	1.43	13	1	1	92.86
	H = 24	84.4	0.9758	0.0028	1.43	13	2	1	86.67
	H = 48	84.25	0.9752	0.0029	1.43	12	2	2	85

Table 3: Performance results of the LPV model and the PWARX model for river level forecasts and for lead times 6h, 12h, 24h and 48h.



Figure 8: (a) Observed level (solid black line) together with PWA model output (solid red line) (b) Prediction error, with H = 1 and f = 24

Figures 9, 10 and 11 show long-term forecasts of the three most important 586 flood events during the validation period, respectively during December,  $13^{th}$ , 587 2017, November 19<sup>th</sup>, 2016 and January 3<sup>rd</sup>, 2016, with the smaller lead 588 time H = 6 and the larger lead time H = 48 in order to assess the error 589 propagation due to the multiple iterations. One can note that the PWARX 590 output deviation does not significantly increase with the prediction horizon. 591 This is due to the high accuracy of the estimated model during the training 592 thus becoming the main condition for PWARX precise forecasts in the case of 593 an exact knowledge of the rainfall forecasts. Errors in water level forecasting 594 in long term will hence be attributed to the errors in the rainfall forecasts 595 which not surprisingly are higher in longer term. 596

597

We also note a worse deviation for multiple hydrographs *i.e.* several

<sup>598</sup> successive flood peaks like the phenomenon observed during November  $19^{th}$ , <sup>599</sup> 2016. It is most likely due to the fact that the data set used for training <sup>600</sup> the model does not probably include the exact similar event, which is the <sup>601</sup> common shortcoming of the black box approaches.



Figure 9: Observed river level (black line) together with PWA short-term forecasts f = 6 h (blue line) and long-term forecasts f = 48 h (red line) of the flood of the December,  $13^{th}$ , 2017. The blue crosses indicate the time when forecasting is done.

#### 602 5. Conclusion

This paper deals with the stream level prediction of a river using a hybrid model over a short-term and long-term prediction horizon in order to prevent damages by a short-term forecast of a flood. The choice of a nonlinear model is justified by numerous factors among them the soil saturation caused notably by intense rainy events during the previous days or the evapotranspiration of the soil and the canopy water interception. The choice of



Figure 10: Observed river level (black line) together with PWA short-term forecasts f = 6 h (blue line) and long-term forecasts f = 48 h (red line) of the flood of November  $19^{th}$ , 2016. The blue crosses indicate the time when forecasting is done.

a piecewise linear structure as a hybrid system is justified by the fact that a 609 river acts differently for the same amount of precipitation and this can be as-610 similated to a switch between multiple models that might be linear according 611 to the presented results, allowing the usage of an easy implementable algo-612 rithm with a low processing time according to the number of treated data 613 samples. In addition, the utilization of an unsupervised clustering technique 614 for the PWARX model identification facilitates the task of achieving a num-615 ber of modes with no *a priori* knowledge and non systematic search. This 616 is possible thanks to the algorithm based on Dempster-Shafer adapted for 617 PWARX modeling. The main difficulty is to fix optimally the value of the 618 main tuning parameter consisting in the number of the nearest neighbors. 619 If this is properly done, one can achieve high performances both in a short-620



Figure 11: Observed river level (black line) together with PWA short-term forecasts f = 6 h (blue line) and long-term forecasts f = 48 h (red line) of the flood of January  $3^{rd}$ , 2016. The blue crosses indicate the time when forecasting is done.

term and a long-term flood forecasting in terms of peak value, the instant 621 when the threshold is crossed and the instant of the peak occurrence. It is 622 however important to mention that the rainfall measurement issued by the 623 rain gauges does not necessarily translate the real precipitation spilled on the 624 catchment. This information has to be cross-checked with the weather radar 625 information in order to reduce the uncertainty on any rainfall-runoff model 626 input. This paper presented the first step of an exploratory work on the 627 usage of PWARX systems in the flood forecast field. Future works will con-628 cern the robustness of these models to an uncertainty on the rainfall which 629 has to be replaced by its forecast, a sensitivity analysis on the parameters, 630 the model calibration and validation on a larger data base and finally, the 631 evaluation of the model use in real-life settings. 632

### References

- C. Perrin, V. Andréassian, Does a large number of parameters enhance model performance ? comparative assessment of common catchment model structures on 429 catchments, Journal of Hydrology 242 (2001) 275–301.
- [2] A. Elshorbagy, G. Corzo, S. Srinivasulu, D. P. Solomatine, Experimental investigation of the predictive capabilities of data driven modeling techniques in hydrology - part 2: Application, Hydrology and Earth System Sciences 14 (2010) 1943 – 1961.
- [3] T. Asefa, M. Kemblowski, M. McKee, A. Khalil, Multi-time scale stream flow predictions: The support vector machines approach, Journal of Hydrology 318 (2006) 7 – 16.
- [4] L. K. A. Siou, A. Johannet, S. Pistre, V. Borrell, Flash floods forecasting in a karstic basin using neural networks: the case of the lez basin (south of france), International Symposium on Karst – ISKA, Malaga. In Advances in research in karst media. Andreo et al eds Springer (2010).
- [5] P. Nayak, K. Sudheer, D. Rangan, K. Ramasastri, A neuro-fuzzy computing technique for modeling hydrological time series, Journal of Hydrology 291 (2004) 52 66. doi:https://doi.org/10.1016/j.jhydrol. 2003.12.010.
- [6] H. Badrzadeh, R. Sarukkalige, A. W. Jayawardena, Intermittent stream flow forecasting and modelling with hybrid wavelet neuro-fuzzy model, Hydrology Research 49 (2017) 27–40. doi:10.2166/nh.2017.163.

- [7] A. B. Dariane, S. Azimi, Streamflow forecasting by combining neural networks and fuzzy models using advanced methods of input variable selection, Journal of Hydroinformatics 20 (2017) 520–532. doi:10.2166/ hydro.2017.076.
- [8] T. K. Chang, A. Talei, C. Quek, V. R. Pauwels, Rainfall-runoff modelling using a self-reliant fuzzy inference network with flexible structure, Journal of Hydrology 564 (2018) 1179 - 1193. doi:https://doi.org/ 10.1016/j.jhydrol.2018.07.074.
- [9] P. C. Young, Time series methods and recursive estimation in hydrological systems analysis, In River flow modelling and forecasting, D. A. Kraijenhoff, J. R. Moll (eds). D. Reidel : Dordrecht (1986) 129 – 180.
- [10] R. Tóth, F. Felici, P. S. C. Heuberger, P. M. J. V. den Hof, Discrete time lpv i/o and state space representations, differences of behavior and pitfalls of interpolation, Proc. of the European Control Conf., Kos, Greece (2007) 5418 – 5425.
- [11] F. Previdi, M. Lovera, Identification of parametrically-varying models for the rainfall-runoff relationship in urban drainage networks, IFAC Proceedings Volumes 42 (2009) 1768 – 1773. 15th IFAC Symposium on System Identification.
- [12] V. Laurain, M. Gilson, S. Payraudeau, C. Grégoire, H. Garnier, Identification de modèles LPV : Application à la modélisation pluie/débit d'un bassin versant viticole, Sixième Conférence Internationale Francophone d'Automatique, CIFA (2010).

- [13] E. Duviella, L. Bako, Predictive black-box modeling approaches for flow forecasting of the liane river., SYSID'12, Bruxelles, Belgium, (2012).
- [14] L. Ljung, System identification : theory for the user (2nd Edition), Prentice Hall PTR, NJ, USA, Upper Saddle River, 1999.
- [15] J. E. Nash, J. V. Sutcliffe, River flow forecasting through conceptual models part I: a discussion of principles, Journal of Hydrology 10 (1970) 282–290.
- [16] C. Perrin, C. Michel, V. Andréassian, Improvement of a parsimonious model for streamflow simulation, Journal of Hydrology 279 (2003) 275 – 289.
- [17] Edijatno, C. Michel, Un modèle pluie-débit journalier à trois paramètres, La Houille Blanche 2 (1989) 113 – 121.
- [18] Edijatno, N. D. O. Nascimento, X. Yang, Z. Makhlouf, C. Michel, GR3J: a daily watershed model with three free parameters, Hydrological Sciences Journal 44 (1999) 263–277.
- [19] F. Bourgin, M. Ramos, G. Thirel, V. Andréassian, Investigating the interactions between data assimilation and post-processing in hydrological ensemble forecasting, Journal of Hydrology 519 (2014) 2775 – 2784. doi:https://doi.org/10.1016/j.jhydrol.2014.07.054.
- [20] A. Ficchi, An adaptive hydrological model for multiple time-steps : diagnostics and improvements based on fluxes consistency, Ph.D. thesis, Université Pierre et Marie Curie - Paris VI, 2017.

- [21] H. Dakhlaoui, D. Ruelland, Y. Tramblay, Z. Bargaoui, Evaluating the robustness of conceptual rainfall-runoff models under climate variability in northern tunisia, Journal of Hydrology 550 (2017) 201 – 217.
- [22] G. Bastin, L. Moens, P. Dierick, Online river flow forecasting with hydromax : successes and challenges after twelve years of experience, In proceedings of the 15th IFAC Symposium on System Identification, Saint-Malo, France, July 6-8 (2009).
- [23] J. Sjoberg, Q. Zhang, L. Ljung, A. Benveniste, B. Delyon, P.-Y. Glorennec, H. Hjalmarsson, A. Juditsky, Nonlinear black-box modeling in system identification: a unified overview, Automatica 31 (1995) 1691 – 1724. Trends in System Identification.
- [24] S. Paoletti, A. L. Juloski, G. Ferrari-Trecate, R. Vidal, Identification of hybrid systems a tutorial, European Journal of Control 13 (2007) 242 – 260.
- [25] R. Vidal, S. Soatto, Yi Ma, S. Sastry, An algebraic geometric approach to the identification of a class of linear hybrid systems, in: 42nd IEEE International Conference on Decision and Control (IEEE Cat. No.03CH37475), volume 1, 2003, pp. 167–172 Vol.1.
- [26] R. Vidal, Identification of PWARX hybrid models with unknown and possibly different orders, in: American Control Conference, volume 1, 2004, pp. 547–552 vol.1.
- [27] A. L. Juloski, S. Weiland, W. P. M. H. Heemels, A bayesian approach

to identification of hybrid systems, IEEE Transactions on Automatic Control 50 (2005) 1520–1533.

- [28] A. Bemporad, A. Garulli, S. Paoletti, A. Vicino, A bounded-error approach to piecewise affine system identification, IEEE Transactions on Automatic Control 50 (2005) 1567–1580. doi:10.1109/TAC.2005. 856667.
- [29] G. Ferrari-Trecate, M. Muselli, D. Liberati, M. Morari, A clustering technique for the identification of piecewise affine systems, Automatica 39(2) (2003) 205–2017.
- [30] F. Lauer, G. Bloch, Piecewise smooth system identification in reproducing kernel hilbert space, in: 53rd IEEE Annual Conference on Decision and Control, 2014, pp. 6498–6503.
- [31] K. Boukharouba, L. Bako, S. Lecoeuche, Identification of piecewise affine systems based on Dempster-Shafer theory, in: IFAC Symposium on System Identification, Saint-Malo, France, 2009, pp. 1662–1667.
- [32] F. Lauer, G. Bloch, R. Vidal, A continuous optimization framework for hybrid system identification, Automatica 47 (2011) 608 - 613. doi:https://doi.org/10.1016/j.automatica.2011.01.020.
- [33] L. Bako, Subspace clustering through parametric representation and sparse optimization, IEEE Signal Processing Letters 21 (2014) 356–360.
- [34] S. Kersting, M. Buss, Recursive estimation in piecewise affine systems using parameter identifiers and concurrent learning, International Journal of Control 0 (2017) 1–18.

- [35] L. Bako, K. Boukharouba, E. Duviella, S. Lecoeuche, A recursive identification algorithm for switched linear/affine models, Nonlinear Analysis: Hybrid Systems 5 (2011) 242 – 253.
- [36] V. Breschi, A. Bemporad, D. Piga, Identification of hybrid and linear parameter varying models via recursive piecewise affine regression and discrimination, in: 2016 European Control Conference (ECC), Aalborg, Denmark, 2016, pp. 2632 – 2637.
- [37] V. Breschi, D. Piga, A. Bemporad, Piecewise affine regression via recursive multiple least squares and multicategory discrimination, Automatica 73 (2016) 155 – 162.
- [38] G. Shafer, A mathematical theory of evidence, Princeton University Press, Princeton, N.J., 1976.
- [39] T. Denoeux, A k-nearest neighbor classification rule based on dempstershafer theory, IEEE Transactions on Systems, Man, and Cybernetics 25 (1995) 804–813. doi:10.1109/21.376493.
- [40] P. Smets, R. Kennes, The transferable belief model, Artificial Intelligence 66 (1994) 191–234.
- [41] B. Hadid, S. Lecoeuche, Data assignment and parameter adaptation for switched LPV system estimation, in: 20th IFAC World Congress, Toulouse, France, 2017, pp. 4564 – 4569.
- [42] V. Vapnik, The nature of statistical learning theory, Springer-Verlag New York Inc, 1995.

- [43] E. J. Bredensteiner, K. P. Bennett, Multicategory classification by support vector machines, Computational Optimization and Applications 12 (1999) 53–79.
- [44] H. Ohlsson, L. Ljung, Identification of switched linear regression models using sum-of-norms regularization, Automatica 49 (2013) 1045 – 1050.
- [45] http://www.hydro.eaufrance.fr (2019).
- [46] D. Norbiato, M. Borga, S. D. Esposti, E. Gaume, S. Anquetin, Flash flood warning based on rainfall thresholds and soil moisture conditions: An assessment for gauged and ungauged basins, Journal of Hydrology 362 (2008) 274 – 290.
- [47] X. Kong, X. Liu, R. Shi, K. Y. Lee, Wind speed prediction using reduced support vector machines with feature selection, Neurocomputing 169 (2015) 449 456. doi:https://doi.org/10.1016/j.neucom.2014.09.
  090, learning for Visual Semantic Understanding in Big Data ESANN 2014 Industrial Data Processing and Analysis.
- [48] A. Hay, The derivation of global estimates from a confusion matrix, International Journal of Remote Sensing 9 (1988) 1395–1398. doi:10. 1080/01431168808954945.
- [49] C. Sammut, G. I. Webb, Encyclopedia of Machine Learning, Springer New York, 2011.
- [50] P. Young, Top-down and data-based mechanistic modelling of rainfallflow dynamics at the catchment scale, Hydrological Processes 17 (2003) 2195 – 2217.

- [51] K. Levenberg, A method for the solution of certain non-linear problems in least squares, Quarterly of Applied Mathematics 2 (1944) 164 - 168. doi:https://doi.org/10.1090/qam/10666.
- [52] D. W. Marquardt, An algorithm for least-squares estimation of nonlinear parameters, Journal of the Society for Industrial and Applied Mathematics 11 (1963) 431-441. URL: http://www.jstor.org/ stable/2098941.
- [53] P. Young, A. Jakeman, Refined instrumental variable methods of recursive time-series analysis part i. single input, single output systems, International Journal of Control 29 (1979) 1–30. doi:https: //doi.org/10.1080/00207177908922676.