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CO-CLUSTERING OF EVOLVING COUNT MATRICES IN PHARMACOVIGILANCE WITH THE DYNAMIC LATENT BLOCK MODEL

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Abstract

Pharmacovigilance is a central medical discipline aiming at monitoring and detecting public health events caused by medicines and vaccines. The purpose of this work is to analyze the notifications of adverse drug reactions (ADRs) gathered by the Regional Center of Pharmacovigilance of Nice (France) between 2010 to 2020. As the current expert detection of safety signals is unfortunately incomplete due to the workload it represents, we investigate here an automatized method of safety signal detection from ADRs data. To this end, we introduce a generative co-clustering model, named dynamic latent block model (dLBM), which extends the classical binary latent block model to the case of dynamic count data. The continuous time is handled by partitioning the considered time period, allowing the detection of temporal breaks in the signals. A SEM-Gibbs algorithm is proposed for inference and the ICL criterion is used for model selection. The application to a large-scale ADRs dataset pointed out that dLBM was not only able to identify clusters that are coherent with retrospective knowledges, in particular for major drug-related crises, but also to detect atypical behaviors, which the health professionals were unaware. Thus, dLBM demonstrated its potential as a routine tool in pharmacovigilance.

Keywords: Co-clustering, pharmacovigilance, latent block model, dynamic count matrices, SEM-Gibbs algorithm

1 Introduction

The World Health Organization defined pharmacovigilance as the science and activities relating to the detection, assessment, understanding and prevention of adverse events or any other drug-related problem (WHO, 2002). The monitoring activities of pharmacovigilance persist even after the drug is on the market to ensure that it remains safe for patients and to assess if previously undetected adverse drug reactions (ADRs) occur. An ADR can be defined as any unwanted or unexpected reaction experienced by an individual after taking a medication (e.g. drug, vaccine, etc.). From the point of view of the public health authorities, it is crucial to report ADRs in order to improve information on medicines and allow possible public health measures, such as withdrawal of a drug, restrictions on indications or interactions between medical products.

As mentioned before, one of the missions of the Regional Centers of Pharmacovigilance (RCPVs) is safety signal detection. However, the method currently used, i.e. manual expert detection of safety signals by the RCPV, despite being unavoidable, has the disadvantage of being incomplete due to its workload. This is why, developing automatized method of safety signal detection is currently a major issue in pharmacovigilance. In such a context, clustering may play an important role in summarizing the information carried out by pharmacovigilance data and identifying patterns of interest. It would be indeed of interest to both cluster the drugs and the adverse reactions to help medical experts in their tasks. The class of methods that allows to perform a simultaneous clustering of both rows (here medical products) and columns (here ADRs) in matrices is known in the statistical literature as co-clustering.

Model-based clustering and co-clustering (Bouveyron et al., 2019) is widely used in an increasing number of different fields. The problem considered in this paper is similar to other situations in which clustering and co-clustering are employed, such as image analysis, pattern recognition, text mining, bioinformatics, information research and social network analysis. In fact, in many applications, it is nowadays frequent to have to summarize large data matrices that evolve in time. For instance, the growing number of users in platforms like Tripadvisor, Amazon, eBay, led the major e-commerce companies to manage massive amounts of data to be able to offer the most suitable products to their customers.

1.1 Pharmacovigilance and the RCPV data

The main mission of the Regional Center of Pharmacovigilance (RCPV), located in the University Hospital of Nice (France), is to gather the notifications of ADRs of three French departments, covering an area of over 2.3 million inhabitants. Such ADR reporting is required by law for every health professional, particularly if serious and/or unexpected, in the respect of the anonymity of both patients and notifiers. Patients, their representatives or patients' associations can also report ADR. Individual patient data collected by the RCPV includes sociodemographic characteristics (age, sex, notifier's department of residence), and the reported ADR (suspected and concomitant drugs, notified effect, date of occurrence, seriousness). The RCPV missions also include answering to questions of health professionals and patients about drugs, promoting a proper use of medicinal product and detecting safety signals about drugs. In France, pharmacovigilance of commercialized drugs is driven by the French Pharmacovigilance network¹, which consists of a ring of 31 RCPV scattered throughout the French territory under the supervision of the national drug agency (ANSM). The French Pharmacovigilance database (FPDB) centralizes all spontaneous reports of adverse drug reactions (ADRs) from the French Pharmacovigilance network, in order to accrue the European database EudraVigilance and the global database Vigibase.

As an illustration of the data we will consider in this work, Figure 1 shows the data structure in three different periods. Each panel shows the interactions between a few molecules (on the x-axis) and a few ADRs (on the y-axis) in the selected period. The distribution of notifications over time is shown in the histograms below, where the considered time period (month) is highlighted in pink. As we can observe on this small part of the whole data set, one clear big event hides all the other patterns that are however of interest.

1.2 Model-based co-clustering

A co-clustering problem can be seen as the simultaneous clustering of rows (individuals) and columns (variables) of a data matrix. To avoid any doubt, it may be useful to point out here the difference in the concepts between co-clustering and bi-clustering. Bi-clustering methods allow to have overlapping clusters and to leave some of the observations unclustered, while in co-clustering each entry belongs to one and only one cluster (it is not possible

¹<https://www.rfcrpv.fr/la-pharmacovigilance/>

to leave some entries unassigned). In the last decades, different co-clustering techniques have been developed for a wide range of applications. Due to the spreading of e-commerce companies (e.g. Amazon, Ebay) and the exponential growth of social networks (e.g. Facebook, LinkedIn, Twitter), application of co-clustering techniques to recommending systems and collaborative filtering have emerged (George and Merugu, 2005; Deodhar and Ghosh, 2010; Xu et al., 2012). Other applications include gene expression analysis (Cheng et al., 2008; Hanisch et al., 2002), text mining and topic modeling (Bergé et al., 2019; Dhillon et al., 2003a; Wang et al., 2009).

Co-clustering methods can be distinguished into metric based approaches such as: non-negative matrix tri-factorization (NMTF) (Labioud and Nadif, 2011; Ding et al., 2006), spectral co-clustering (Dhillon, 2001), information theory (Dhillon et al., 2003b) and model-based co-clustering approaches (Bouveyron et al., 2019).

Among those approaches, model-based co-clustering is widely appreciated for its sound statistical foundations and its flexibility in terms of sparsity and data types. The much popular latent block model (LBM) (Govaert and Nadif, 2003) was introduced for the co-clustering of binary data matrices, based on the assumption that rows and columns are grouped in hidden clusters and that observations within a block (intersection of a row cluster and a column cluster) are independent and identically distributed. Whereas the original formulation of the model dealt with binary data, the model has been extended in the last two decades to counting data (Govaert and Nadif, 2010), continuous data (Lomet, 2012), categorical data (Keribin et al., 2015), ordinal data (Jacques and Biernacki, 2018; Corneli et al., 2020), functional data (Bouveyron et al., 2018) and textual data (Bergé et al., 2019). Several inference procedures have been also proposed for LBM, including likelihood based methods (Govaert and Nadif, 2008), variational inference (Keribin et al., 2012), Bayesian inference (Keribin et al., 2012; Wyse and Friel, 2012) and greedy search approaches (Wyse and Friel, 2012; Wyse et al., 2017). Recently, Boutalbi et al. (2020) proposed the tensor latent block model (TLBM) for co-clustering, whose aim is to simultaneously cluster rows and columns of a 3D matrix, where covariates represent the third dimension. TLBM was also implemented for different types of datasets: continuous data (Gaussian TLBM), binary data (Bernoulli TLBM) and contingency tables (Poisson TLBM).

Among the dynamic extensions of the co-clustering approaches, we can cite the evolutionary spectral co-clustering (ESCC, Green et al., 2011). In that

paper, two different ways of taking into account the historical relationship between the instances and features are proposed, one respect to current (RTC) and the other respect to historical (RTH). The first one gives most importance to the present, considering only the previous time-step, while the second one takes into account all the previous time-steps. For the continuous time modeling, an original approach was proposed by Corneli et al. (2018) for the dynamic stochastic block model, in a slightly different context: the clustering of dynamic networks. The authors considered the interactions between nodes in a dynamic network, where the goal is to cluster both nodes and time intervals into groups. The number of interactions between two nodes is modeled via a non-homogeneous Poisson process (NHPP) whose instantaneous intensity functions are assumed to be constant on each time cluster. A greedy maximization of the exact-ICL (Biernacki et al., 2000; Côme and Latouche, 2015) was used to cluster the nodes, the time intervals and to select the optimal number of clusters.

Concerning co-clustering applications to pharmacovigilance, a seminal article was proposed by Robert et al. (2015). In that paper, the authors introduced the multiple latent block model (MLBM) by extending the latent block model (Govaert and Nadif, 2008) through the construction of one row partition and two columns partitions that respectively rely on two binary matrices containing the relation between the individual and the medical product and the relation between the individual and the ADR.

1.3 Contributions of this work

We consider count data matrices evolving over a time period $[0, T]$, whose number of rows and columns is fixed. The goal of the dynamic latent block model (dLBM), that we propose here as a tool for pharmacovigilance, is to cluster both the rows and the columns of such matrices. We assume that the number of interactions occurring in time between rows and columns follows a non-homogeneous Poisson process (NHPP). To handle the dynamic framework, we led a segmentation over the continuous time such that we obtain as many static matrices as the number of the identified time intervals. As inference procedure, we use the SEM-Gibbs algorithm while the ICL criterion is adopted for selecting the optimal number of clusters. Thus, dLBM provides a meaningful summary of massive datasets, possibly with a large number of missing data. To demonstrate the interest in pharmacovigilance, we run a large-scale experiment on an ADRs dataset from Regional Center

of Pharmacovigilance (RCPV), located in the University Hospital of Nice (France). The interest of this application lies not only in summarizing the massive amount of drug adverse reactions (ADRs) data but also in identifying possible unexpected phenomena, such as atypical side effects of certain types of drugs.

1.4 Organization of the paper

This paper is organized as follows. Section 2 first recalls the original latent block model, before introducing the generative model of dLBM. In Section 3, the inference procedure is detailed and a model selection criterion is discussed. Section 4 presents various experiments on simulated data to assess the model performances. In Section 5, an application on a real ADRs dataset is presented to illustrate the potential of dLBM in pharmacovigilance. Section 6 provides some concluding remarks.

2 The dynamic latent block model

In this section, we introduce the dynamic latent block model (dLBM). The main goal of this model is the simultaneous clustering of rows and columns of high-dimensional sparse matrices in a dynamic time framework. The data we consider are organized such that the rows (drugs in pharmacovigilance application) are indexed by $i = 1, \dots, N$ and the columns (adversarial effects) by $j = 1, \dots, P$. Moreover, we consider a fixed time period $[0, T]$ during which the total number of rows, N , and columns, P , is fixed. We indicate as $\mathcal{X}(t)$ the $N \times P$ matrix that contains the number of interactions occurring between the individual i and the item j at time $t \in [0, T]$. The first part of this section reviews the latent block model, while, in the second part, the proposed dynamic extension is introduced.

2.1 The latent block model

Let us first neglect the time dimension and recall the original latent block model (Govaert and Nadif, 2010), assuming that the time period is restricted to a single time point. The data structure is therefore a $N \times P$ random matrix $X = \{X_{ij}\}_{i \in 1, \dots, N, j \in 1, \dots, P}$. Rows and columns of X are assumed to be clustered respectively into K and L groups, such that the data belonging to the same block are independent and identically distributed. More formally, the latent structure of rows and columns of X is identified by the following latent variables:

- $Z := \{z_{ik}\}_{i \in 1, \dots, N, k \in 1, \dots, K}$ represents the clustering of rows into K groups: $\mathcal{A}_1, \dots, \mathcal{A}_K$. The row i belongs to cluster \mathcal{A}_k iff $z_{ik} = 1$;
- $W := \{w_{j\ell}\}_{j \in 1, \dots, P, \ell \in 1, \dots, L}$ represents the clustering of columns into L groups: $\mathcal{B}_1, \dots, \mathcal{B}_L$. The column j belongs to cluster \mathcal{B}_ℓ iff $w_{j\ell} = 1$.

Moreover, Z and W are assumed to be independent and distributed according to multinomial distributions:

$$p(Z|\gamma) = \prod_{k=1}^K \gamma_k^{|\mathcal{A}_k|}, \quad p(W|\rho) = \prod_{\ell=1}^L \rho_\ell^{|\mathcal{B}_\ell|},$$

where $\gamma_k = \mathbb{P}\{z_{ik} = 1\}$, $\rho_\ell = \mathbb{P}\{w_{j\ell} = 1\}$, $\sum_{k=1}^K \gamma_k = 1$, $\sum_{\ell=1}^L \rho_\ell = 1$, and $|\mathcal{A}_k|$ and $|\mathcal{B}_\ell|$ respectively represent the number of rows in cluster \mathcal{A}_k and the number of columns in cluster \mathcal{B}_ℓ .

The LBM model further assumes the entries X_{ij} are independent, conditionally to Z and W , and their distribution $\varphi(\cdot, \alpha)$ belongs to the same parametric family, where the parameter α only depends on the given block:

$$X_{ij} \mid z_{ik}w_{j\ell} = 1 \sim \varphi(X_{ij}, \alpha_{k\ell}). \quad (1)$$

With these assumptions the complete data likelihood can be written as:

$$p(X, Z, W; \theta) = p(Z; \theta)p(W; \theta)p(X|Z, W; \theta) = \prod_{k=1}^K \gamma_k^{|\mathcal{A}_k|} \prod_{\ell=1}^L \rho_\ell^{|\mathcal{B}_\ell|} \prod_{i,k} \prod_{j,\ell} \varphi(X_{ij}; \alpha_{k\ell})^{z_{ik}w_{j\ell}}. \quad (2)$$

As mentioned in Section 1.2 various versions of LBM have been proposed along the decades. The first one was proposed for the binary case, where $\varphi(X_{ij}, \alpha_{k\ell})$ is a Bernoulli distribution with $p(X_{ij} = 1|Z, W; \theta) = \alpha_{k\ell}$. Since this article focuses on count data, the most suitable extension of LBM for our purpose is the Poisson LBM. In fact, this model is based on the assumption that, conditionally on Z and W , the entries X_{ij} follow a Poisson distribution $\mathcal{P}(\lambda_{k\ell})$, i.e:

$$\varphi(X_{ij}; \lambda_{k\ell}) = \left(\frac{\lambda_{k\ell}^{X_{ij}}}{X_{ij}!} \exp(-\lambda_{k\ell}) \right). \quad (3)$$

2.2 Modeling the dynamic framework

Let us now introduce the time dimension such that \mathcal{X}_{ij} is time dependent. Thus, $\mathcal{X}_{ij}(t)$, $t \in [0, T]$, represents the cumulative number of interactions at time t between i and j .

A possible approach for the dynamic modeling relies on non-homogeneous Poisson processes (NHPPs), thus assuming that $\{\mathcal{X}_{ij}(\cdot)\}_{i,j}$ are independent point processes, with instantaneous intensity functions $\lambda_{ij}(t)$:

$$\mathcal{X}_{ij}(t) \sim \mathcal{P} \left(\int_0^t \lambda_{ij}(u) du \right), \quad (4)$$

where $\mathcal{P}(\lambda)$ denotes the Poisson probability mass function of parameter λ . With the notation adopted so far, we thus assume the existence of $N \times P$ independent Poisson processes.

In order to cluster both the rows and the columns, we further assume that the intensity function $\lambda_{ij}(t)$ only depends on the respective clusters of row i and column j :

$$\mathcal{X}_{ij}(t) \mid z_{ik}, w_{j\ell} = 1 \sim \mathcal{P} \left(\int_0^t \lambda_{k\ell}(u) du \right).$$

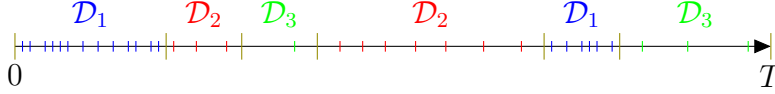


Figure 2: Time clusters.

For further use, let us introduce the parameter $\lambda := (\lambda_{k\ell}(t))_{k \leq K; \ell \leq L}$. Given the above assumptions, the conditional distribution of the number of interactions between i and j , over the time interval $[s, t]$, where $0 \leq s \leq t \leq T$, is:

$$p(\mathcal{X}_{ij}(t) - \mathcal{X}_{ij}(s) \mid z_{ik}w_{j\ell} = 1, \lambda) = \frac{(\int_s^t \lambda_{k\ell}(v) dv)^{\mathcal{X}_{ij}(t) - \mathcal{X}_{ij}(s)}}{(\mathcal{X}_{ij}(t) - \mathcal{X}_{ij}(s))!} \exp\left(-\int_s^t \lambda_{k\ell}(v) dv\right). \quad (5)$$

2.2.1 A discrete time version

In order to ease the understanding of the dynamic model and to make the inference tractable, we also operate a clustering over the time dimension. Let us first introduce a discretization of the considered time interval $[0, T]$. Thus, without loss of generality the following partition of $[0, T]$ is introduced:

$$0 = t_0 < t_1 < \dots < t_U = T, \quad (6)$$

where the U intervals, $I_u = [t_{u-1}, t_u[$, will also be clustered. The number of interactions between i and j on the time interval I_u can be therefore summarized by X_{iju} :

$$X_{iju} := \mathcal{X}_{ij}(t_u) - \mathcal{X}_{ij}(t_{u-1}), \quad \forall(i, j, u),$$

where $\mathcal{X}_{ij}(t_u)$ represents the cumulative number of interactions at time t_u between i and j . Hence, we introduce the tensor $X := \{X_{iju}\}_{iju}$ with dimension $N \times P \times U$ that contains the number of interactions between any observation and feature pair at any given time interval. We can also see X as a time series (along the third dimension) of incidence matrices.

Since our goal is to perform clustering over the time dimension as well, each time interval I_1, \dots, I_U is also assumed to be assigned to a hidden time cluster $\mathcal{D}_1, \dots, \mathcal{D}_c$. To model the membership to time clusters, a new latent variable S has to be introduced, such that $s_u = c$ if the time interval I_u belongs to the time cluster \mathcal{D}_c . As it is shown in Figure 2, it is worth noticing

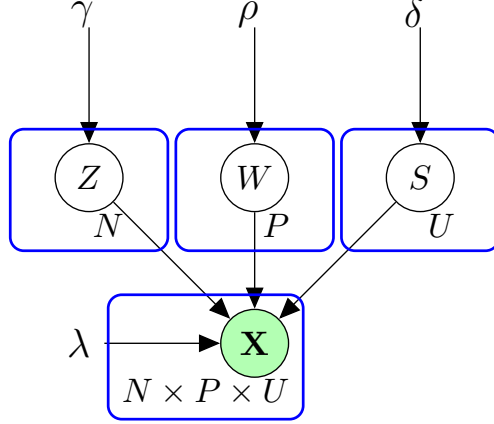


Figure 3: Graphical representation of dLBM.

that a specific time cluster can occur more than once in the temporal line when a similar interactivity pattern is repeated in time. Furthermore, as for Z and W , we assume that S follows a multinomial distribution:

$$p(S | \delta) = \prod_{c=1}^C \delta_c^{|\mathcal{D}_c|}, \quad (7)$$

where $\delta_c = \mathbb{P}\{s_{uc} = 1\}$; $\sum_{c=1}^C \delta_c = 1$ and $|\mathcal{D}_c|$ represents the number of time intervals in the cluster \mathcal{D}_c .

Once these additional assumptions have been made, we can rewrite Eq. (4) considering that the intensity functions are stepwise constant on each time cluster \mathcal{D}_c . Thus:

$$X_{iju} | z_{ik} w_{jl} s_{uc} = 1 \sim \mathcal{P}(\lambda_{klc} \Delta_u), \quad (8)$$

where Δ_u indicates the length of the interval I_u . Henceforth, in order to simplify the exposition, we assume that Δ_u is constant, $\Delta_u = \Delta$. We can finally set $\Delta = 1$ without loss of generality. A graphical representation of dLBM can be seen in Figure 3.

From Eqs. (3)-(8), it holds that:

$$p(X_{iju} | z_{ik} w_{jl} s_{uc} = 1, \lambda_{klc}) = \left(\frac{(\lambda_{klc})^{X_{iju}}}{X_{iju}!} \exp(-\lambda_{klc}) \right). \quad (9)$$

Therefore, we can introduce the $K \times L \times C$ tensor λ , whose elements are denoted by $\lambda_{k\ell c}$.

It is now possible to write the complete data likelihood of the model:

$$p(X, Z, W, S | \gamma, \rho, \delta, \lambda) = p(Z | \gamma) p(W | \rho) p(S | \delta) p(X | Z, W, S, \lambda), \quad (10)$$

where $p(Z | \gamma)$, $p(W | \rho)$ and $p(S | \delta)$ were defined in the previous section. The conditional distribution of X , given Z , W , and S , can be easily obtained from Eq. (9) by independence:

$$p(X | Z, W, S, \lambda) = \prod_{k,\ell,c} \left(\frac{(\lambda_{k\ell c})^{R_{k\ell c}}}{P_{k\ell c}} \exp(-|\mathcal{A}_k| |\mathcal{B}_\ell| |\mathcal{D}_c| \lambda_{k\ell c}) \right), \quad (11)$$

where $R_{k\ell c} = \sum_{i=1}^N \sum_{j=1}^P \sum_{u=1}^U z_{ik} w_{j\ell} s_{uc} X_{iju}$ and $P_{k\ell c} = \prod_{i=1}^N \prod_{j=1}^P \prod_{u=1}^U (z_{ik} w_{j\ell} s_{uc} X_{iju})!$.

Denoting by θ the set of all model parameters, $\theta = (\gamma, \rho, \delta, \lambda)$, the log-likelihood can be finally written as:

$$\ell(\theta; X) = \sum_Z \sum_W \sum_S \log p(X, Z, W, S | \theta). \quad (12)$$

2.3 Link with related models

At this point, dLBM can be related with the following models:

- If we do not take into account the time dependency, assuming that the time period is restricted to a single time point t_0 , dLBM coincides with the Poisson LBM.
- dLBM reduces to dSBM (Corneli et al., 2016) if the row individuals are the same as the column. In that case, in fact, Z would be equal to W and, then consequently, $K = L$. Therefore, in this way, we pass from an incidence matrix to an adjacency matrix.
- If $C = U$, dLBM corresponds to TensorLBM (Boutalbi et al., 2020), where the third dimension is considered but the slices are not clustered. In fact, when the contingency table case is analyzed, the authors consider a Poisson TensorLBM, where for each slice a , the entries are distributed according to a $Poisson(\lambda_{ij}^a)$.

3 Inference algorithm and model selection

3.1 Model inference

For the co-clustering model based approach outlined in the previous section, as well as for standard mixture models, a direct maximization of the log-likelihood with respect to the model parameters is not feasible. A widely used approach to overcome this problem is the expectation-maximization (EM) algorithm (Dempster et al., 1977). It consists in alternating two steps, E step and M step, in order to maximize a lower bound of the log-likelihood. The EM-algorithm exploits the following decomposition of the log-likelihood:

$$\ell(\theta; X) = \log p(X | \theta) = \mathcal{L}(q(Z, W, S); \theta) + KL(q(Z, W, S) || p(Z, W, S | X, \theta)), \quad (13)$$

with:

$$\mathcal{L}(q(Z, W, S); \theta) = \sum_{Z, W, S} q(Z, W, S) \log \frac{p(X, Z, W, S | \theta)}{q(Z, W, S)}, \quad (14)$$

where $\mathcal{L}(\cdot; \theta)$ represents a lower bound of the log-likelihood while $KL(q||p)$ is the Kullback-Leibler divergence between q and p . During the E-Step, the algorithm maximizes the lower bound $\mathcal{L}(q(Z, W, S); \theta)$ with respect to $q(\cdot)$ for a given value of θ . From Eq.(13), we can assess that maximizing $\mathcal{L}(q; \theta)$ is equivalent to minimizing $KL(q||p(\cdot|X, \theta))$ and since the Kullback-Leibler divergence cannot be negative, we conclude that the lower bound is maximized when $KL(q||p(\cdot|X, \theta)) = 0$. This leads us to the conclusion that $\mathcal{L}(q(Z, W, S); \theta)$ is optimized when $q^*(Z, W, S) = p(Z, W, S | X, \theta)$. The M-Step usually consists in maximizing $\mathcal{L}(q(Z, W, S); \theta)$ over θ , keeping $q = q^*$ fixed, in order to obtain an updated version of the parameter, θ .

However, in the context of the present work the joint posterior distribution $p(Z, W, S | X, \theta)$ is not computationally tractable as well. To go through this limitation, we propose to approximate it through a Gibbs sampler within the E-step. Such an approach was proposed by Keribin et al. (2010) and exploited, for instance, by Bouveyron et al. (2018) for the functional latent block model (funLBM). The resulting stochastic alternative of the EM algorithm, called SEM-Gibbs algorithm, starts with some initial values of the parameter $\theta^{(0)}$, the column clusters $W^{(0)}$ and the time clusters $S^{(0)}$. In this way, at the h^{th} iteration the algorithm alternates the following SE step and M step:

SE step: a partition for Z , W and S is drawn according to $q^*(Z, W, S) = p(Z, W, S | X, \theta)$, which is approximated by making use of a Gibbs sampler, using the current values of the parameter set θ . In this way, the unknown labels are simulated from their posterior distribution, given the observed data and the parameter set. It consists in executing a small number of iterations of the following three steps:

- 1) Generate the row partition $z_i^{(h+1)} = (z_{i1}^{(h+1)}, \dots, z_{iK}^{(h+1)}) | X, W^{(h)}, S^{(h)}$ according to $z_i^{(h+1)} \sim \mathcal{M}(1, \tilde{z}_{i1}, \dots, \tilde{z}_{iK})$, for all $1 \leq i \leq N$ and $1 \leq k \leq K$, where:

$$\tilde{z}_{ik} = p(z_{ik} = 1 | X, W^{(h)}, S^{(h)}; \theta^{(h)}) = \frac{\gamma_k^{(h)} f_k(X_i | W^{(h)}, S^{(h)}; \theta^{(h)})}{\sum_{k'} \gamma_{k'}^{(h)} f_{k'}(X_i | W^{(h)}, S^{(h)}; \theta^{(h)})},$$

where $X_i = (X_{iju})_{ju}$ and $f_k(X_i | W^{(h)}, S^{(h)}; \theta^{(h)}) = \prod_{j\ell} \prod_{uc} p(X_{iju}; \theta_{k\ell c}^{(h)}) w_{j\ell}^{(h)} s_{uc}^{(h)}$.

- 2) Generate the column partition $w_j^{(h+1)} = (w_{j1}^{(h+1)}, \dots, w_{jL}^{(h+1)}) | X, Z^{(h+1)}, S^{(h)}$ according to $w_j^{(h+1)} \sim \mathcal{M}(1, \tilde{w}_{j1}, \dots, \tilde{w}_{jL})$, for all $1 \leq j \leq P$ and $1 \leq \ell \leq L$, where:

$$\tilde{w}_{j\ell} = p(w_{j\ell} = 1 | X, Z^{(h+1)}, S^{(h)}; \theta^{(h)}) = \frac{\rho_\ell^{(h)} f_\ell(X_j | Z^{(h+1)}, S^{(h)}; \theta^{(h)})}{\sum_{\ell'} \rho_{\ell'} f_{\ell'}(X_j | Z^{(h+1)}, S^{(h)}; \theta^{(h)})},$$

where $X_j = (X_{iju})_{iu}$ and $f_\ell(X_j | Z^{(h+1)}, S^{(h)}; \theta^{(h)}) = \prod_{ik} \prod_{uc} p(X; \theta_{k\ell c}^{(h)}) z_{ik}^{(h+1)} s_{uc}^{(h)}$.

- 3) Generate the time cluster partition $s_u^{(h+1)} = (s_{u1}^{(h+1)}, \dots, s_{uC}^{(h+1)}) | X, Z^{(h+1)}, W^{(h+1)}$ according to $s_u^{(h+1)} \sim \mathcal{M}(1, \tilde{s}_{u1}, \dots, \tilde{s}_{uC})$, for all $1 \leq u \leq U$ and $1 \leq c \leq C$, where:

$$\tilde{s}_{uc} = p(s_{uc} = 1 | X, Z^{(h+1)}, W^{(h+1)}; \theta^{(h)}) = \frac{\delta_c^{(h)} f_c(X_u | Z^{(h+1)}, W^{(h+1)}; \theta^{(h)})}{\sum_{c'} \delta_{c'} f_{c'}(X_u | Z^{(h+1)}, W^{(h+1)}; \theta^{(h)})},$$

where $X_u = (X_{iju})_{ij}$ and $f_c(X_u | Z^{(h+1)}, W^{(h+1)}; \theta^{(h)}) = \prod_{ik} \prod_{j\ell} p(X_{iju}; \theta_{k\ell c}^{(h)}) z_{ik}^{(h+1)} w_{j\ell}^{(h+1)}$.

M step: in this step, $\mathcal{L}(q^*(Z, W, S); \theta^{(h)})$ is maximized with respect to θ , where:

$$\begin{aligned}\mathcal{L}(q^*(Z, W, S); \theta^{(h)}) &\simeq \sum_{Z, W, S} p(Z, W, S | X, \theta^{(h)}) \log \frac{p(X, Z, W, S | \theta)}{p(Z, W, S | \theta^{(h)})} \\ &\simeq E[\log(p(X, Z^{(h+1)}, W^{(h+1)}, S^{(h+1)} | \theta) | \theta^{(h)})] + \xi,\end{aligned}$$

where ξ is a constant term related to θ . This conditional expectation of the complete data log-likelihood can be written in a developed form as follows:

$$\begin{aligned}E[\log(p(X, Z^{(h+1)}, W^{(h+1)}, S^{(h+1)} | \theta) | \theta^{(h)})] &= \sum_{i,k} z_{ik}^{(h+1)} \log \gamma_k + \sum_{j,\ell} w_{j\ell}^{(h+1)} \log \rho_\ell + \\ &+ \sum_{u,c} s_{u,c}^{(h+1)} \log \delta_c + \sum_{i,k} \sum_{j,\ell} \sum_{u,c} z_{ik}^{(h+1)} w_{j\ell}^{(h+1)} s_{uc}^{(h+1)} \log(p(X_{iju} | \theta_{k\ell c})).\end{aligned}\tag{15}$$

The last term of the previous equation can be further developed as:

$$\begin{aligned}&\sum_{i,k} \sum_{j,\ell} \sum_{u,c} z_{ik}^{(h+1)} w_{j\ell}^{(h+1)} s_{uc}^{(h+1)} \log(p(X_{iju} | \theta_{k\ell c})) = \\ &\sum_{i,k} \sum_{j,\ell} \sum_{u,c} z_{ik}^{(h+1)} w_{j\ell}^{(h+1)} s_{uc}^{(h+1)} \log \left[\frac{(\lambda_{k\ell c}^{(h)})^{X_{iju}}}{X_{iju}!} \exp(-\lambda_{k\ell c}^{(h)}) \right] = \\ &\sum_{i,k} \sum_{j,\ell} \sum_{u,c} z_{ik}^{(h+1)} w_{j\ell}^{(h+1)} s_{uc}^{(h+1)} \left[X_{iju} \log(\lambda_{k\ell c}^{(h)}) - \log(X_{iju}!) - \lambda_{k\ell c}^{(h)} \right].\end{aligned}\tag{16}$$

Thanks to the previous equation, the parameter set $\theta^{(h+1)}$ can be estimated. The mixture proportions are updated as follows (proof in Appendix B):

$$\gamma_k^{(h+1)} = \frac{1}{N} \sum_i z_{ik}^{(h+1)}, \quad \rho_\ell^{(h+1)} = \frac{1}{P} \sum_j w_{j\ell}^{(h+1)}, \quad \delta_c^{(h+1)} = \frac{1}{U} \sum_u s_{uc}^{(h+1)}.$$

Moreover, the ML estimator of $\lambda_{k\ell c}$ is defined as follows (proof in Appendix C):

$$\lambda_{k\ell c}^{(h+1)} = \frac{R_{k\ell c}^{(h+1)}}{|\mathcal{A}_k| |\mathcal{B}_\ell| |\mathcal{D}_c|}, \quad \forall (k, \ell, c).$$

The algorithm runs for a certain number of iterations of the two steps and, after the burn-in period, we can obtain the final estimation of the parameter set $\hat{\theta}$ by computing the mean from the sampled distribution. Finally, the optimal values for Z , W and S are estimated by the mode of their sampled distributions.

3.2 Model selection

Up to now, we have assumed that the number of row clusters (K), column clusters (L) and time clusters (C) was known. However, for real data sets, this assumption is of course unrealistic. For this reason, our purpose in this section is to define a model selection criterion that can automatically identify the optimal number of clusters that are appropriate for the data at hand. The model selection approach is considered. We propose to rely on ICL (Integrated Completed Likelihood, Biernacki et al. (2000)) to approximate the complete-data integrated log-likelihood. We derived the formulation of the ICL criterion for the model proposed above:

$$\begin{aligned}
 ICL(K, L, C) = \log p(X, \hat{Z}, \hat{W}, \hat{S}; \hat{\theta}) - \frac{K-1}{2} \log N + \\
 - \frac{L-1}{2} \log P - \frac{C-1}{2} \log U - \frac{KLC}{2} \log(NPU)
 \end{aligned} \tag{17}$$

The triplet $(\hat{K}, \hat{L}, \hat{C})$ that leads to the highest value for the ICL is considered as the most meaningful for those data.

4 Numerical experiments

The main purpose of this section is to highlight the most important features of dLBM over simulated datasets. We aim at demonstrating the validity of the inference algorithm and model selection criterion, presented in the previous sections. The first experiment consists in applying dLBM to an easy scenario to explain its main outputs. Then, the second example shows a model selection application on 25 simulated datasets. In the third experiment, we compare the performances of dLBM with some state-of-the-art methods in three simulated scenarios.

4.1 Introductory example

As a first example, we simulate a dataset with $K = 3$ groups of rows, $L = 2$ groups of columns and $C = 2$ groups of time clusters, with a level of sparsity $\tau = 0.97$. Table 1 shows the main features of this dataset.

N	P	U	τ	K	L	C	γ	ρ	δ
200	200	150	0.97	3	2	2	(0.1, 0.35, 0.55)	(0.2, 0.8)	(0.6, 0.4)

Table 1: Parameter values for the first simulated dataset

We fitted dLBM to simulated dataset with the actual values for K , L and C to show the ability of the model to fully recover the model parameters. Figure 4 shows the evolution of the complete data log-likelihood. As we can see the convergence is reached in less than 10 iterations. Figure 5 shows the evolution of the estimated mixture parameters $\hat{\gamma}$, $\hat{\rho}$ and $\hat{\delta}$ along the iterations of the SEM-Gibbs algorithm. Comparing the values reached by each line with the actual values of the model parameters showed in Table 1, we can observe that dLBM fully recovers the original values in few iterations. Moreover, Figure 6 shows a bar plot of the number of interactions between rows and columns of the array X for each time period, where the two different time clusters are identified by different colors. We can easily deduce that dLBM selects the two time clusters in a meaningful way in terms of level of counted interactions in each time cluster.

Figure 7 shows the value of the estimated intensity parameter $\hat{\lambda}$ for each cluster of rows and columns where different colors represent different time clusters. For instance, the algorithm detects that, in the first cluster of rows

and the first cluster of columns (Block(1,1)), there is an high intensity of interactions in both of the time clusters. Figure 8 and Figure 9, display the data structure before running dLBM and the reorganized incidence matrices one for each time cluster. To this end, rows and columns of the incidence matrix are permuted, thanks to the estimates \hat{Z} and \hat{W} , in such a way that nearby rows (columns) belong to the same cluster of rows (columns). The blocks are also delimited by black dashed lines.

Finally, to evaluate the performance of the model in identifying the correct rows, columns and times partitions, we use the adjusted Rand index (ARI) Rand (1971) for all of the three variables. The adjusted Rand index, from a mathematical point of view, is closely related to the accuracy measure, however it is a commonly used method for evaluating clustering problems since it can be applied for measuring the similarity between two partitions even with different number of clusters and it is invariant to label switching. The closer the index is to 1, the more two label vectors are similar to each other. We compared the original matrices Z , W and S , with the estimates \hat{Z} , \hat{W} and \hat{S} given by the output of the dLBM. The model obtained an ARI index of 1 for rows, columns and times partitions. Thus, we can conclude that our algorithm perfectly identifies the composition of the original clusters.

4.2 Model selection experiment

In the previous experiment, we assumed to know the value of K , L and C . In this section we aim at validating the ICL criterion for model selection. To do that, 25 independent datasets are generated with the setup indicated in Table 2. dLBM is applied on those simulated datasets for values of K , L and C ranging from 1 to 6. The results are sorted according to the ICL values. Table 3 shows the percentage of selections by ICL criterion on 25 simulated datasets. The highlighted cell corresponds to the actual value of K , L and C . ICL succeeds 64% of the time to identify the correct model. It is worth to notice that, when ICL does not select the right combination of K , L and C , the wrongly selected models are close to the simulated one. In particular, 28% of the selections only differ from the actual model by one cluster, on one of the three cluster dimensions.

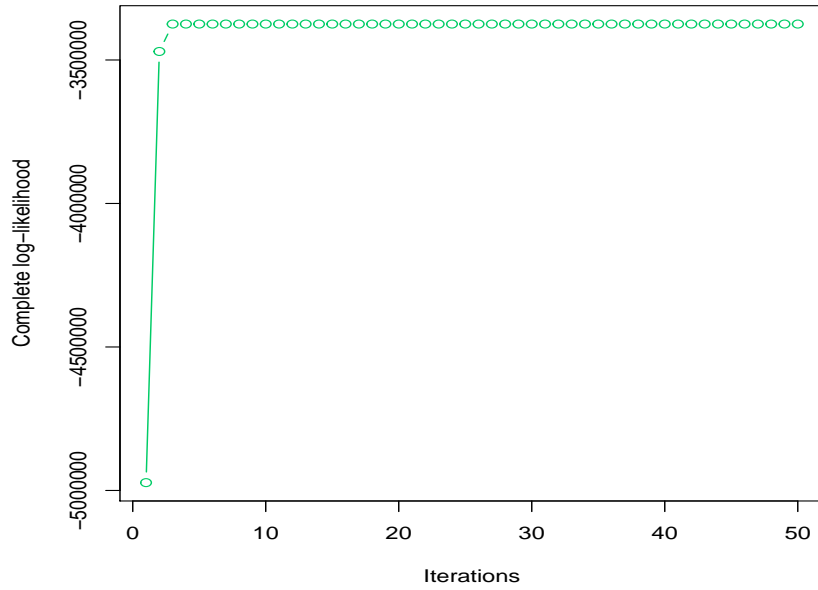


Figure 4: Complete data log-likelihood over the iterations of the dLBM algorithm.

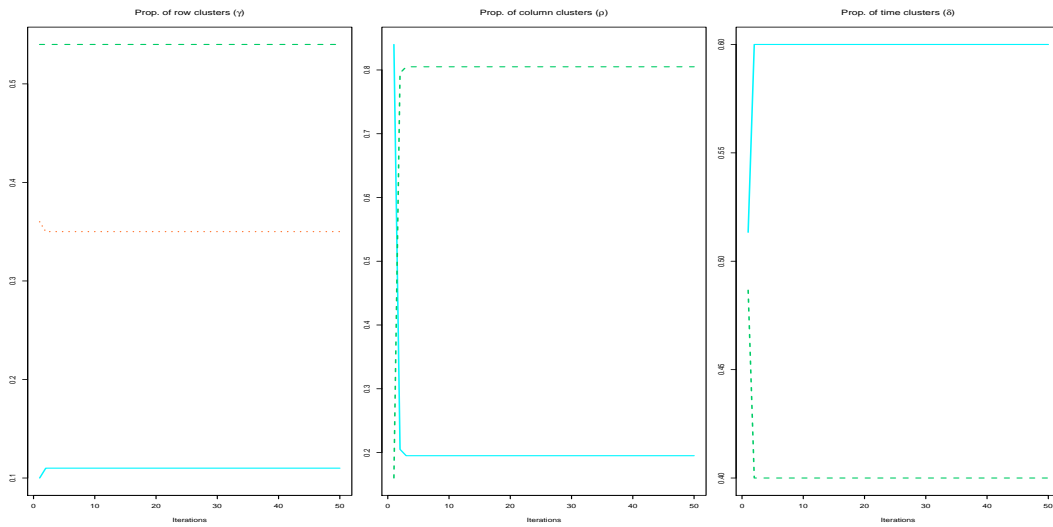


Figure 5: Estimates of the mixture parameters of the first simulated dataset.

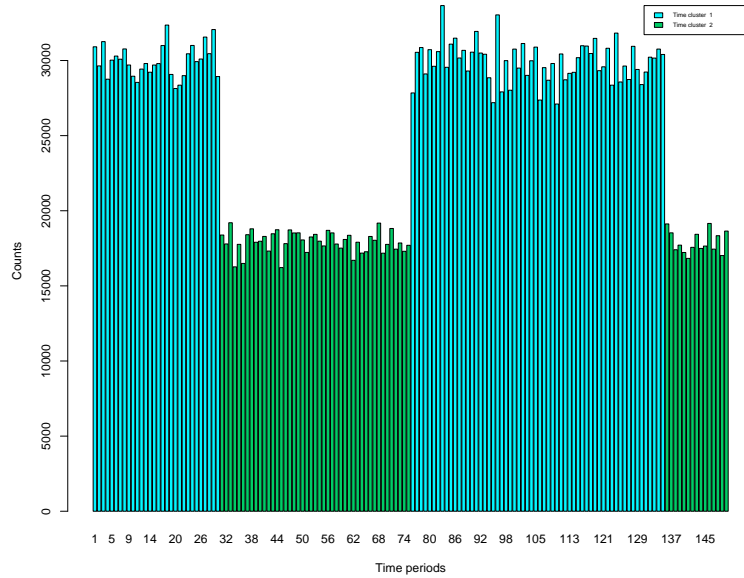


Figure 6: Time periods representation of the first simulated dataset, where different colors correspond to different time clusters.

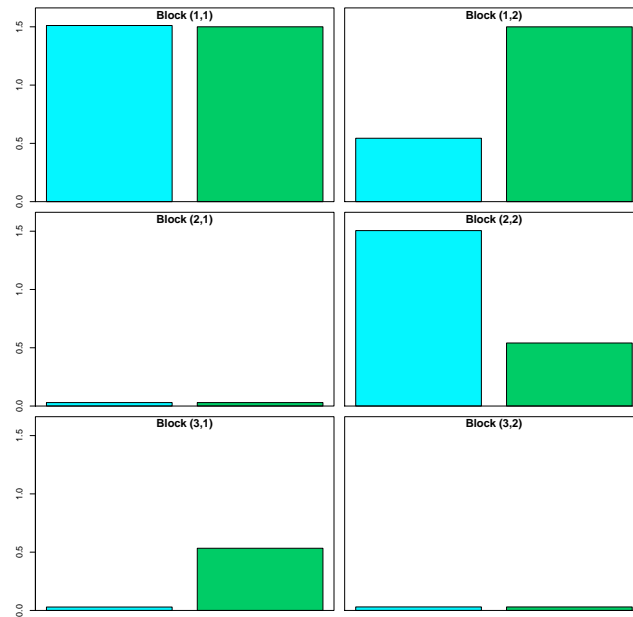


Figure 7: Estimated λ values for each cluster of rows and columns.

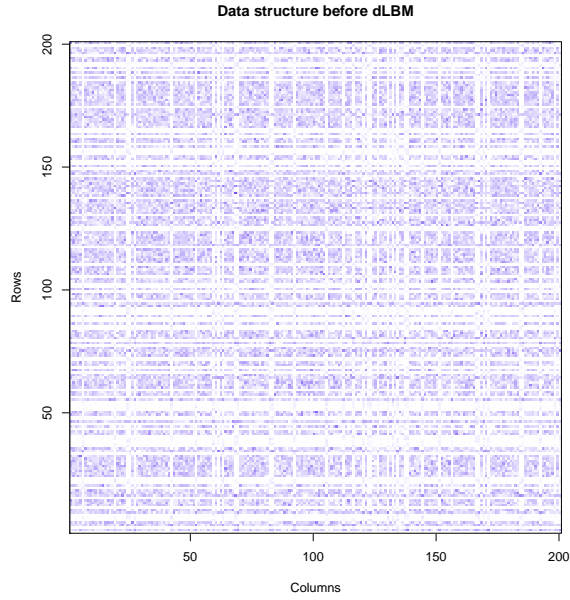
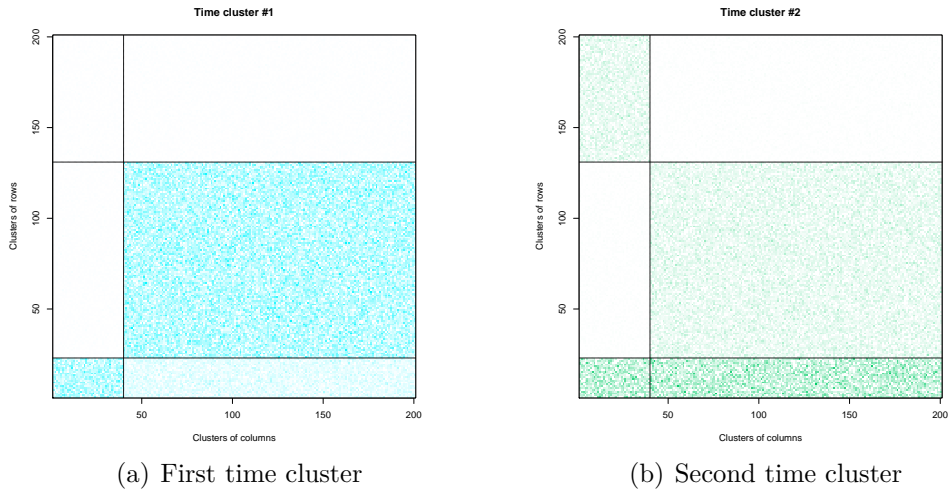


Figure 8: Unordered data structure before running dLBM



(a) First time cluster

(b) Second time cluster

Figure 9: Reorganized incidence matrices, one for each time cluster, according to the estimates \hat{Z} and \hat{W} . Nearby rows (columns) belong to the same cluster of rows (columns). The blocks are also delimited by black dashed lines.

N	P	U	τ	K	L	C	γ	ρ	δ
200	200	200	0.97	4	3	3	(0.2,0.4,0.1,0.3)	(0.4,0.3,0.3)	(0.25,0.3,0.45)

Table 2: Parameter values for the second simulated dataset.

C = 3							C = 4						
K/L	1	2	3	4	5	6	K/L	1	2	3	4	5	6
1	0	0	0	0	0	0	1	0	0	0	0	0	0
2	0	0	0	0	0	0	2	0	0	0	0	0	0
3	0	0	0	0	0	0	3	0	0	0	0	0	0
4	0	0	64	8	0	0	4	0	0	8	0	0	0
5	0	0	12	0	4	0	5	0	0	0	4	0	0
6	0	0	0	0	0	0	6	0	0	0	0	0	0

Table 3: Model selection. Percentage of selections by ICL criterion on 25 simulated datasets. The highlighted cell corresponds to the actual value of K, L and C.

4.3 Benchmark study

The goal of this third experiment is to compare dLBM with some state-of-the-art methods in terms of recovering the data structure. dLBM is compared with TensorLBM (Boutalbi et al., 2020) where, in absence of the original code, we set the number of time clusters of dLBM equal to the number of time intervals, $C = U$, and with the Poisson LBM by making use of the `bikm1` package (Robert et al., 2020). Since LBM supports only two dimensions, we shrink the third dimension summing up alternatively on rows, columns and slices, obtaining respectively the Row_LBM, Col_LBM and Slice_LBM methods.

We chose to evaluate the results with the ARI index by comparing the resulting cluster partitions with the simulated ones. To make this comparison more complete, we defined three simulation scenarios ("Easy", "Medium" and "Hard"), detailed in Table 4. Table 5 displays the results of this comparison, in terms of average ARI values, reported with standard deviations. The dash indicates that no value is reported because the calculation is not allowed by the model.

In the "Easy" situation, dLBM works perfectly. Also TensorLBM provides excellent results, even though calculated only on rows and columns. Row_LBM

and Col_LBM only give good results on one dimension, while Slice_LBM produces extremely low results.

In "Medium" and "Hard" situations, dLBM continues to obtain excellent results, although not perfect, due to the increasing complexity of the proposed situations. The three other LBM models perform poorly, while TensorLBM obtains rather high ARI values. Specifically, in the "Hard" situation the ARI value of TensorLBM on the columns is slightly higher than that of dLBM, even if the one calculated on the rows partitions is lower.

Scenario	N	P	U	K	L	C	γ	ρ	δ	λ
Easy	250	250	100	3	2	2	(0.1,0.35,0.55)	(0.2,0.8)	(0.6,0.4)	Λ_A
Medium				4	3	3	(0.2,0.4,0.1,0.3)	(0.4,0.3,0.3)	(0.25,0.3,0.45)	Λ_B
Hard										Λ_C

Table 4: Parameter values for the three simulation scenarios (see Appendix C for details about Λ_A , Λ_B and Λ_C).

Scenario A - Easy			
	ARI_Rows	ARI_Cols	ARI_Slices
dLBM	1 ± 0	1 ± 0	1 ± 0
TensorLBM	0.8 ± 0.3	1 ± 0	-
Row_LBM	-	0.12 ± 0.21	0.96 ± 0.2
Col_LBM	0.1 ± 0.21	-	0.92 ± 0.2
Slice_LBM	0.09 ± 0.2	0.13 ± 0.22	-
Scenario B - Medium			
	ARI_Rows	ARI_Cols	ARI_Slices
dLBM	0.89 ± 0.17	1 ± 0	1 ± 0
TensorLBM	0.74 ± 0.18	1 ± 0	-
Row_LBM	-	0.13 ± 0.21	0.12 ± 0.23
Col_LBM	0.09 ± 0.21	-	0.15 ± 0.23
Slice_LBM	0.1 ± 0.2	0.14 ± 0.21	-
Scenario C - Hard			
	ARI_Rows	ARI_Cols	ARI_Slices
dLBM	0.79 ± 0.19	0.68 ± 0.22	0.63 ± 0.18
TensorLBM	0.64 ± 0.21	0.71 ± 0.19	-
Row_LBM	-	0.09 ± 0.14	0.09 ± 0.15
Col_LBM	0.12 ± 0.14	-	0.12 ± 0.15
Slice_LBM	0.2 ± 0.18	0.25 ± 0.22	-

Table 5: Co-clustering results for dLBM, TensorLBM, and LBM applied respectively by summing up the rows (Row_LBM), the columns (Col_LBM) and the slices (Slice_LBM) on 25 simulated data according to the three scenarios. Average ARI values are reported with standard deviations.

5 Analysis of the adverse drug reaction dataset

This section focuses now on the application of dLBM to a large-scale pharmacovigilance dataset, with the aim of illustrating the potential of the tool for such studies.

5.1 Protocol and data

This section considers a large dataset consisting of ADR data collected by the Regional Center of Pharmacovigilance (RCPV), located in the University Hospital of Nice (France). The center covers an area of over 2.3 million inhabitants and receives notifications about ADRs from different channels: a website² form that everyone can freely fill and send, phone calls, emails, medical visits at the hospital units, etc. A time horizon of 10 years is considered, from January 1st, 2010 to September 30th, 2020, the unity measure for time intervals is a month ($\Delta_u = \Delta = 1$ month). The overall dataset is made of by 44,269 declarations, for which the market name of the drug, the notified ADR, the channel used for the declaration and its origin, as well as an identification number and the reception date are reported. To prevent the same medicine from being considered more than once if reported under slightly different names, we decided to use the international nonproprietary name (INN) of the drug (to simplify the comprehension, the INNs would be referred as drugs for the rest of the study). Moreover, we only considered molecules and ADRs that were notified more than 10 times over the 10 years. The resulting dataset contains 542 drugs, 586 ADRs and 129 months with 13,363 non-zero entries.

Figure 1 shows a histogram representing the frequency of declarations arrived at the pharmacovigilance center from 2010 to 2020, ordered by month. It can easily be noticed that, during the year 2017, there is an extremely uncommon behavior in the progress of notifications to the pharmacovigilance center. In fact, in that year an unexpected rise of reports for ADRs happened concerning two specific drugs: Mirena[®] and Lévothyrox[®]. Mirena[®] has been available in Europe since 1995. This birth control product contains a hormone called levonorgestrel. In 2017, concerns regarding ADRs associated with the use of levonorgestrel releasing intra-uterine device (IUD) started to grown with a media coverage peak occurred in May 2017, which resulted in a massive wave

²<https://signalement.social-sante.gouv.fr>

of ADRs reports from patients to French RCPVs (Langlade et al., 2019). Also, Lévothyrox[®] has been marketed in France for about 40 years as a treatment for hypothyroidism and, in 2017, a new formula was introduced on the market. The Lévothyrox[®] case had an extremely high media coverage in France: the RCPVs received 18,241 reports of Lévothyrox[®] ADRs in 2017 only. Lévothyrox[®] spontaneous reports represent almost the 90% of all the spontaneous notifications that the Nice center received from patients in 2017. This phenomenon has been fully described in a recent article of Viard et al. (2019).

From Figure 1, one can understand the difficulty to work with such data which contain signals of very different amplitude. Indeed, behind those very visible effects, many ADR signals need to be detected for obvious public health reasons. In particular, those data also contain ADR reports regarding Médiator[®], which is here far less visible than Lévothyrox[®] and Mirena, but also led to many avoidable serious cardiovascular diseases. This is why, we expect dLBM to be a useful tool to reveal such hidden signals.

5.2 Summary of the results

Remembering that our aim is to find an underlying latent structure in our dataset by applying co-clustering on the three dimensions of the array X with dimension $542 \times 586 \times 129$, we have run dLBM for different values of K , L and C . Specifically, we tested rows (here drugs), columns (here ADRs) and times groups ranging from 1 to 12. The ICL criterion identified the optimal values for the triplet $(\hat{K}, \hat{L}, \hat{C})$ as: $\hat{K} = 7$, $\hat{L} = 10$, $\hat{C} = 6$.

Figure 10 shows the frequency of the declarations received by the RCPV from 2010 to 2020, sorted by month, where the colors represent the identified time clusters. Figure 11 displays the estimated intensity function representation. In particular, this figure is very helpful for giving an overview of the relationships between drug clusters and ADR clusters and how they evolve over time. The colors refer to different time clusters and the brighter the color, the stronger the relation (i.e. the expected number of notifications in the time unit) between drug cluster and ADR cluster. Finally, Figure 12 shows more specifically the evolution of the relationship between drug clusters and ADR clusters over time. In fact, each panel represents a cluster of drugs and within them each line identifies a cluster of ADRs and its intensity changes over time.

Time clusters Starting from the analysis of the time clusters, one can easily notice on Figure 10 that the segmentation proposed by the algorithm confirms our knowledge about the previous mentioned health scandals while revealing a time structure more complex than expected. In fact, while cluster 1 and cluster 2 include various time intervals, cluster 3 clearly refers to the health crisis due to the Mirena[®] scandal while cluster 4 relates to the peak period in the Lévothyrox[®] crisis. Time clusters 5 and 6 refer to the final stage of the Lévothyrox[®] crisis, when generics were introduced to the market. It is worth noticing that without the dLBM application it would have been impossible to detect the presence of other health scandal just before the one of Lévothyrox[®]. In fact, looking at Figure 10, one can see that the increase of declarations during the Mirena[®] health crisis are completely masked by the Lévothyrox[®] ones.

Drug clusters The clusters of drugs identified by the algorithm are also coherent with retrospective knowledge and adequately represent the variety of drugs present in the dataset. In particular, cluster 1, cluster 6 and cluster 7 are very specific, with one element only: they correspond respectively to lévothyroxine (Lévothyrox[®] and generics), benfluorex (Médiator[®]) and lévonorgestrel (Mirena[®]). It is worth noticing that Médiator[®]³ was involved (like Lévothyrox[®] and Mirena[®]) in an important health scandal in 2009-2010. Moreover, cluster 2 contains the five most frequently reported drugs and cluster 5 contains other common drugs, while cluster 4 is very large and heterogeneous, with drugs that are rarely reported and finally cluster 3 contains drugs that cause bleeding.

ADR clusters Concerning the clusters of ADRs, cluster 3 (e.g. coma, confusion, hepatic cytolysis, etc) and cluster 8 (e.g. agitation, agranulocytosis, arthralgia, etc.) contain the most frequently notified ADRs. Cluster 1 contains recurring ADRs (e.g. sweats, transient ischemic accident, lactic acidosis, etc.) but less than the other two previously mentioned. Cluster 2 (e.g. anemia, hemorrhagic stroke) and cluster 4 (e.g. hemorrhagic shock, deglobulisation, etc.) respectively include the most and the less frequent bleeding related ADRs. Cluster 7 is composed of ADRs clearly related to Lévothyrox[®] and Mirena[®] (e.g hair loss, cramps, insomnia, etc.). In cluster 10

³[https://www.ansm.sante.fr/Dossiers/Mediator-R/Mediator-R-et-accompagnement-des-personnes/\(offset\)/0](https://www.ansm.sante.fr/Dossiers/Mediator-R/Mediator-R-et-accompagnement-des-personnes/(offset)/0)

there are general ADRs, although it contains some ADRs specifically related to Lévothyrox[®] and Médiator[®] (e.g. respectively abnormal TSH, valvular disease, etc.). Finally, cluster 5, 6 and 9 contain more general and nonspecific ADRs.

5.3 Detailed results

Time clusters Figure 11 is a graphical representation of the estimated intensity functions. It gives a clear idea about the relationships between clusters of molecules and clusters of ADRs, with respect to time clusters. In particular:

- Time cluster 1: here one can notice the presence of all the drug clusters, with different levels of intensity. The peculiarity of this cluster lies in the strong presence of cluster 6 of drugs which gradually disappears in subsequent temporal clusters. In fact, it contains the drug benfluorex (Médiator[®]) which in 2010 was involved in a major health scandal and it has strong interactions with clusters 5, 6 and 10 of ADRs. It is worth noticing that dLBM managed to highlight this peculiarity that cannot be detected by simply looking at Figure 1. In the first time cluster it can also be noticed a strong relation between the cluster 3 of drugs (drugs that cause bleeding) and clusters 2 and 4 of ADRs which is coherent.
- Time cluster 2: the presence of Médiator[®] decreases while the interactions between the drugs that cause bleeding (cluster 3) and clusters 2 and 4 of ADRs is still strong. That cluster represents the actual profile of notifications received by RCPV. Similarly, the second cluster of drugs (the most frequently used) appears to have ADRs in almost all clusters, especially the third. We also notice the presence of Lévothyrox[®] (cluster 1) with ADRs especially in clusters 6 and 7.
- Time cluster 3: this cluster includes two months only: 05-06/2017. This period refers to the Mirena[®] scandal (cluster 7) with ADRs in clusters 5, 6, 7 and 9. They are not very specific but it may suggest a hormonal cause. In this matrix all the drug clusters are present (even though with a lower intensity with respect to Mirena[®]), with the exception of Médiator[®] which is not reappearing subsequently.

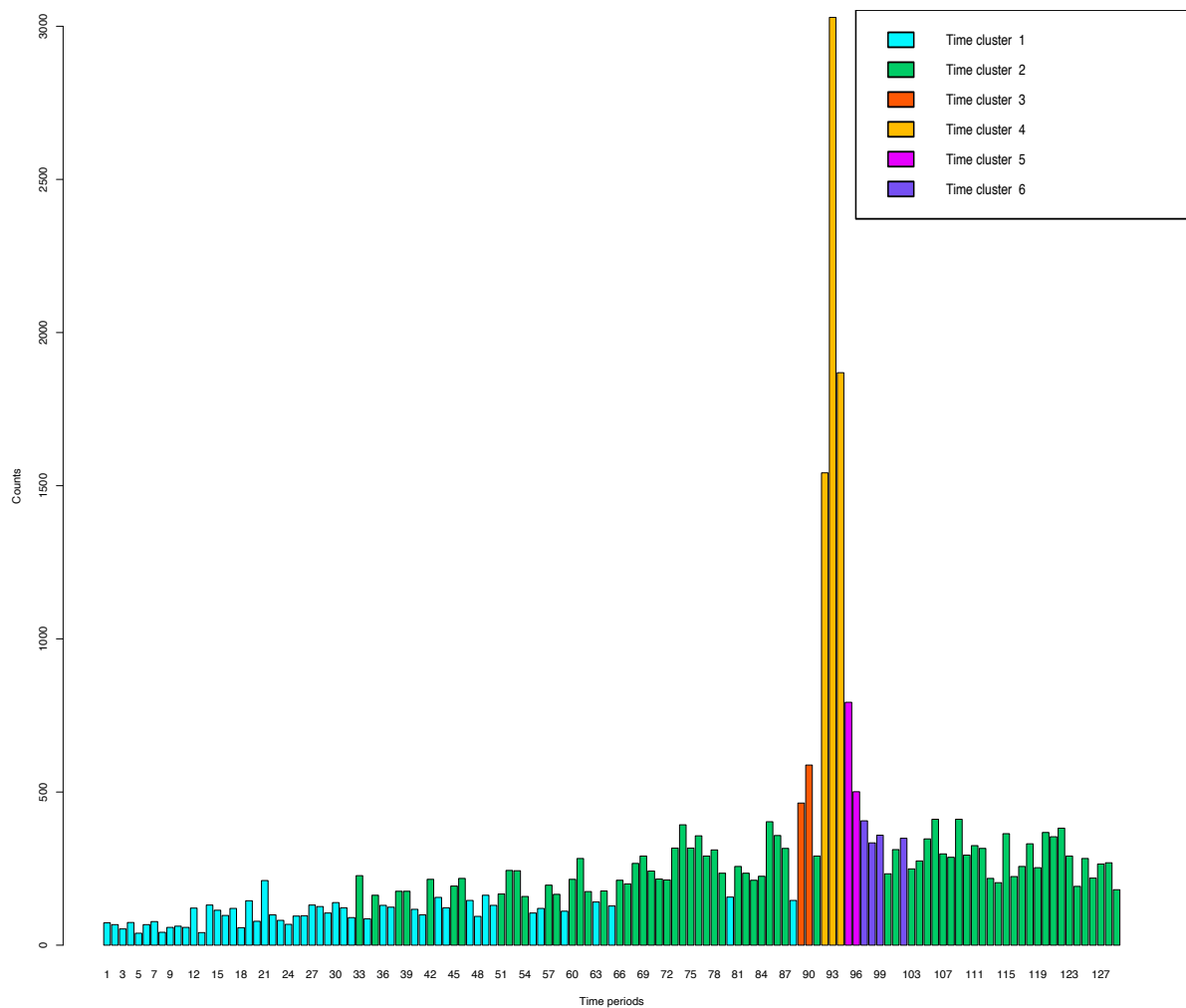


Figure 10: Number of declarations received by the pharmacovigilance center from 2010 to 2020, sorted by month, where the colors represent the time clusters.

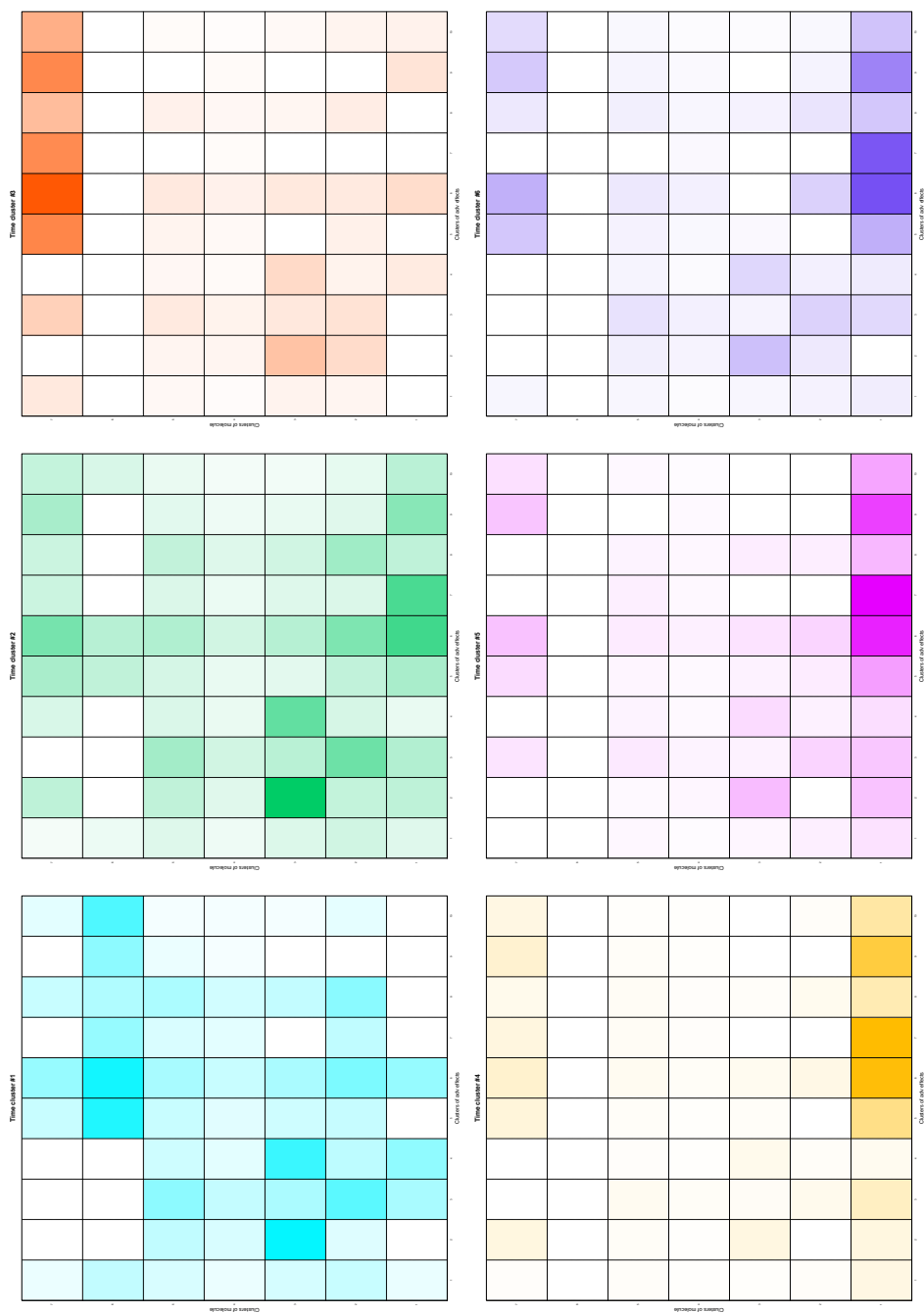


Figure 11: Estimated intensity function representation. Colors refer to different time clusters. The brighter the color, the stronger the relation between rows (drugs) clusters and columns (ADRs) clusters.

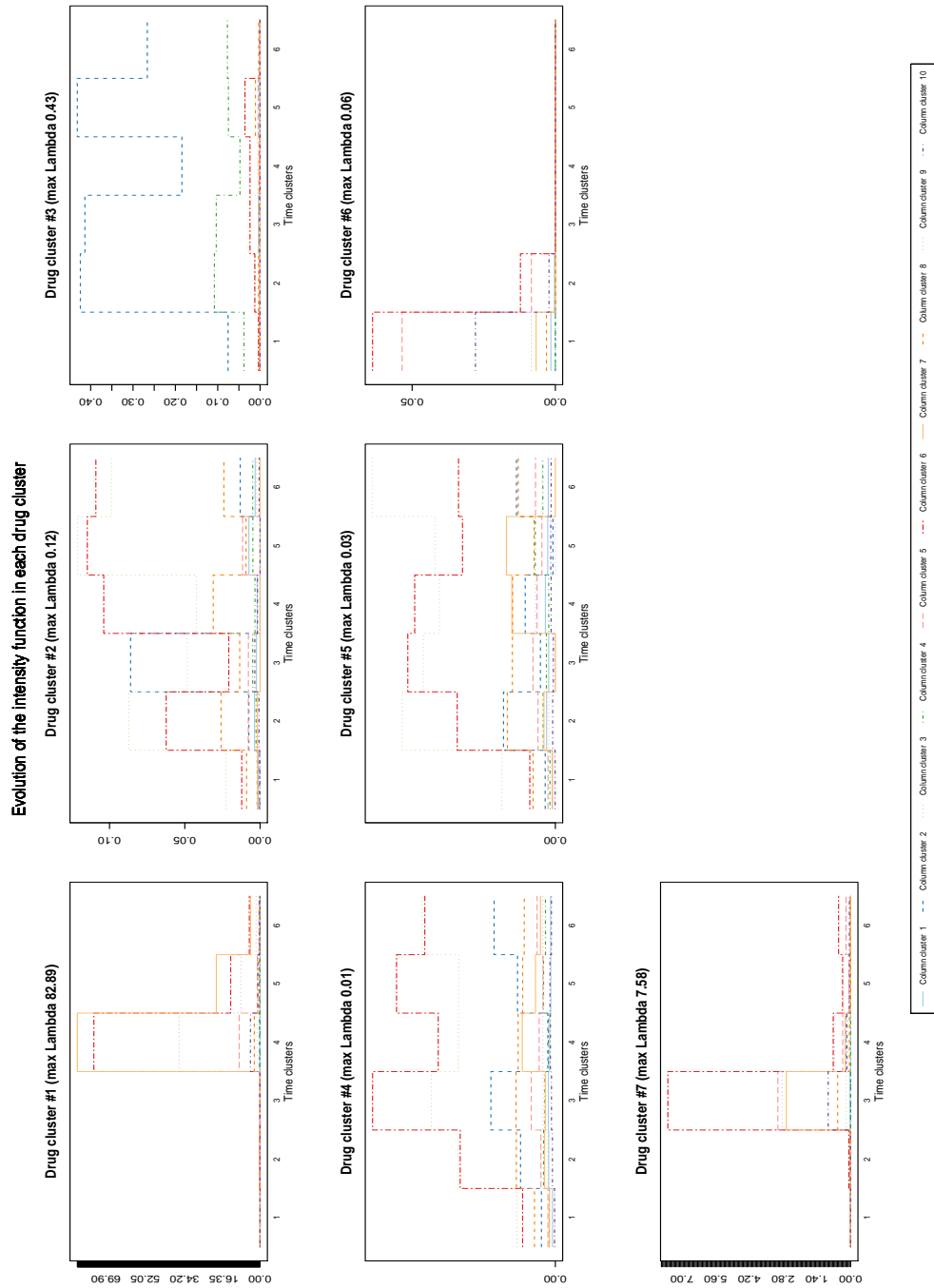


Figure 12: Evolution of the relation between each drug cluster and the all ADR clusters over time. Each color corresponds to a different cluster of adverse drug reaction.

- Time cluster 4: it refers to the Lévothyrox[®] peak going from August to October 2017. Unlike previous clusters where most drugs clusters were present, this cluster only recognizes 2 drugs: Lévothyrox[®] (mainly) and Mirena[®] (weakly). Consequently, the interactions that stand out are those of Lévothyrox[®] with ADRs in clusters 6, 7 and 9. They are not very specific but coherent with the statements received by RCPV.
- Time cluster 5: it refers to 11-12/2017 and it is characterized by a reduction of Lévothyrox[®] declarations. Compared to the previous cluster, we note a reappearance of other drug clusters even if the intensities remain low compared to time cluster 1 and 2. The drugs/ADRs combinations remain those of the previous time cluster.
- Time cluster 6: it refers to the 1st semester 2018, which corresponds to end of Lévothyrox[®] crisis. Globally it is similar to the two previous time clusters with some small variations on the intensities of the drug/ADRs pairs.

Drug clusters For a more in-depth analysis regarding the evolution of drug clusters over time and their interactions with the clusters of ADRs, we can refer to Figure 12. The following remarks derive:

- Drug cluster 1: this cluster refers to Lévothyrox[®] and its generics. There are almost no declared effects during the first three time clusters, from the fourth time cluster we observe a peak of declarations which corresponds to the start of the Lévothyrox[®] crisis, especially for ADRs in cluster 6, 7 and 9 and to a lesser extent on 5. These four clusters recognized by dLBM, include all of the ADRs described during Lévothyrox[®] crisis, namely hormonal ADRs (weight gain), general ADRs (fatigue, cramps) and neuro-psychic ADRs (anxiety, irritability, sleep disturbances). Time cluster 5 marks a decrease in Lévothyrox[®] reports in terms of numbers but the ADRs profile remains similar to the previous time cluster. It should be noticed that generics of Lévothyrox[®] began to be available from mid-October 2017, which could explain this decrease in the number of reports: patients started to have therapeutic alternatives. Finally, time cluster 6 represents the end of the Lévothyrox[®] crisis with a clear decrease in the number of reports.

- Drug cluster 2: this cluster includes drugs that are very frequently prescribed, the ADRs profile is globally constant over time with a predominance of clusters 3, 6 and 8, with variations in terms of proportions according to the time clusters. Cluster 3 of ADRs corresponds to frequent and generally serious ADRs. Cluster 8 also includes ADRs that are generally serious but a little less often reported than cluster 3. Cluster 6 corresponds to general ADRs that can be found with many other drugs (especially Lévothyrox[®]).
- Drug cluster 3: this cluster includes coagulation drugs whose main ADRs are bleeding, hence the predominance of cluster 2 and 4. The application of dLBM led us to identify, from the temporal point of view, 3 interesting events: the increase in ADRs between time clusters 1 and 2, a significant decrease in the number of declared ADRs in time cluster 4 (Lévothyrox[®] crisis) and a marked regression of cluster 2 in time cluster 6 (late Lévothyrox[®]) but without cluster 4 being affected.
- Drug cluster 4: this cluster includes a fairly large set of drugs declared relatively frequently and commonly prescribed, but disparate in terms of their therapeutic uses or their ADRs profile. However, we observe a predominance of ADRs of cluster 3 and 6 (general ADRs). At the temporal level, we observe an overall decrease in ADRs in time cluster 4 (Lévothyrox[®] crisis).
- Drug cluster 5: as for drug cluster 4, this cluster includes many heterogeneous drugs. The ADRs profile is similar to that of drug cluster 4, which is coherent with the fact that these two clusters are similar.
- Drug cluster 6: this is the Médiator[®] cluster, the ADRs are concentrated in time clusters 1 and 2 with a decrease in the number of ADRs in the second one. This is coherent with the history of the drug (Médiator[®] scandal happened in 2009 with withdrawal of the market). Regarding the profile of ADRs, we can notice a predominance of clusters 6, 5, and 10.
- Drug cluster 7: this is the Mirena[®] cluster, ADRs declarations predominate in time cluster 3 (Mirena[®] crisis) then drop drastically. At the level of ADRs profile, cluster 6 predominates (general ADRs), then come cluster 5, 7 and 9.

5.4 Discussion

In this application to pharmacovigilance, dLBM proved to be a very useful tool for identifying phenomena that would have been difficult to detect otherwise, even by an expert eye. In fact, dLBM revealed that in addition to Lévothyrox[®] health crisis, which was the one with the widest media coverage, two other major events have occurred. The first one concerning Médiator[®], which took place in 2009-2010, and the second one concerning Mirena[®], which took place in the first half of 2017. In addition, dLBM was also able to put in light some unexpected variations of notifications such as an under-notification of bleeding related ADRs during Lévothyrox[®] crisis. Bleeding related ADRs were expected to be constant over time because of the follow-up made the RCPV to monitor ADRs of direct oral anticoagulants (DOAs), a recent class of anticoagulant. However, the Lévothyrox[®] crisis has caused such an overload of work that the DOAs follow-up have been temporarily interrupted. Another thing that dLBM has highlighted is the existence of 3 different phases during the Lévothyrox[®] crisis corresponding to the reporting peak, the marketing period of generics and the end of the crisis, respectively. Those phases were not noticed by the RCPV staff during the Lévothyrox[®] crisis. In general, we can conclude that dLBM could be extremely useful as a routine tool for signal detection, since it might help health professionals to identify structural changes or patterns of interest and, perhaps, prevent some of the consequences a health crisis can lead to.

6 Conclusion

This work is born out of the need to analyze and summarize the adverse drug reaction data collected by the Regional Center of Pharmacovigilance (RCPV), located in the University Hospital of Nice (France), from 2010 to 2020. We have proposed a dynamic co-clustering technique, with the purpose of simultaneously performing clustering of rows (drugs), columns (adversarial effects) and slices (time dimension). We have proposed to consider a dynamic framework because it is of great interest to look for structural changes in the way existing drug/ADR clusters (uncovered, too) interact with each other. To this end, we have introduced a generative model, named dynamic latent block model (dLBM). The dynamic time modeling relies on non-homogeneous Poisson processes, with a latent partition of time intervals. Inference is done using a SEM-Gibbs algorithm and the ICL criterion is used for model selection. dLBM was fitted to a large-scale data set supplied by the RCPV of Nice. In this context, dLBM provided meaningful segmentations of drugs, adverse drug reactions and time. Its potential use by medical authorities for identifying meaningful pharmacovigilance patterns looks very promising.

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A Estimation of the mixture proportions

The proof about how to obtain the updated mixture proportions is only shown for the estimation of parameter $\gamma_k^{(h+1)}$ because for the estimation of the other parameters, ρ and δ , the procedure is similar:

$$\begin{aligned}
 p(Z|\gamma) &= \mathcal{L}(\gamma; Z) = N! \prod_{i=1}^N \prod_{k=1}^K \frac{\gamma_k}{z_{ik}!}; \\
 \ell(\gamma_k; z_{ik}^{(h+1)}) &= \log \mathcal{L}(\gamma_k, z_{ik}^{(h+1)}) = \log \left(N! \prod_{i=1}^N \prod_{k=1}^K \frac{\gamma_k}{z_{ik}^{(h+1)}!} \right) = \\
 &= \log N! + \sum_{i=1}^N \sum_{k=1}^K z_{ik}^{(h+1)} \log \gamma_k - \sum_{i=1}^N \sum_{k=1}^K \log z_{ik}^{(h+1)}!
 \end{aligned}$$

For a constrained maximization of this quantity we employ the Lagrange Multipliers, taking into account the constraint $\sum_{k=1}^K \gamma_k = 1$.

$$\begin{aligned}
 \mathcal{L}(\gamma_k; \lambda) &= \ell(\gamma_k; z_{ik}^{(h+1)}) + \lambda \left(1 - \sum_{k=1}^K \gamma_k \right) \\
 \frac{\partial \mathcal{L}(\gamma_k; \lambda)}{\partial \gamma_k} &= \frac{\partial \ell(\gamma_k; z_{ik}^{(h+1)})}{\partial \gamma_k} + \frac{\partial \lambda (1 - \sum_k \gamma_k)}{\partial \gamma_k} = 0 \\
 \frac{\partial \sum_{i=1}^N \sum_{k=1}^K z_{ik}^{(h+1)} \log \gamma_k}{\partial \gamma_k} - \lambda \frac{\partial \sum_{k=1}^K \gamma_k}{\partial \gamma_k} &= 0 \\
 \frac{\sum_{i=1}^N z_{ik}^{(h+1)}}{\gamma_k} - \lambda &= 0 \\
 \sum_{i=1}^N z_{ik}^{(h+1)} &= \lambda \gamma_k \Rightarrow \frac{\sum_{i=1}^N z_{ik}^{(h+1)}}{\lambda} = \gamma_k
 \end{aligned}$$

Since λ is equal to N :

$$\sum_{k=1}^K \sum_{i=1}^N \frac{z_{ik}^{(h+1)}}{\lambda} = \sum_{k=1}^K \gamma_k \Rightarrow \frac{1}{\lambda} \sum_{k=1}^K \sum_{i=1}^N z_{ik}^{(h+1)} = 1;$$

we can conclude that the estimation of $\gamma_k^{(h+1)}$ is the following:

$$\gamma_k^{(h+1)} = \frac{1}{N} \sum_{i=1}^N z_{ik}^{(h+1)}$$

B Maximum likelihood estimator of $\lambda_{k\ell c}$

The maximum likelihood estimator of $\lambda_{k\ell c}$ is obtained through the following process:

$$\log L(\lambda|X, Z, W, S) = \sum_{k=1}^K \sum_{\ell=1}^L \sum_{c=1}^C (R_{k\ell c} \log \lambda_{k\ell c} - |\mathcal{A}_k||\mathcal{B}_\ell||\mathcal{D}_c| \lambda_{k\ell c} + c)$$

where c is a constant that includes all the terms that does not depend on λ .

$$\frac{\partial \log \mathcal{L}(\lambda|X, Z, W, S)}{\partial \lambda} = \frac{R_{k\ell c}}{\lambda_{k\ell c}} - |\mathcal{A}_k||\mathcal{B}_\ell||\mathcal{D}_c| = 0 \Rightarrow \hat{\lambda}_{k\ell c} = \frac{R_{k\ell c}}{|\mathcal{A}_k||\mathcal{B}_\ell||\mathcal{D}_c|}$$

C Intensity functions in the three scenarios

From Table 4, the scenarios "Easy" and "Medium" may look the same. However, the main difference between the two scenarios is the value assumed by the intensity function λ . The values of this parameter in the three different scenarios are:

- Scenario A - Easy: $\lambda = \Lambda_A$

$$\Lambda_A[:, 1] = \begin{bmatrix} 50 & 18 \\ 1 & 1 \\ 1 & 50 \end{bmatrix}; \Lambda_A[:, 2] = \begin{bmatrix} 50 & 50 \\ 18 & 1 \\ 1 & 18 \end{bmatrix}$$

- Scenario B - Medium: $\lambda = \Lambda_B$

$$\Lambda_B[:, 1] = \begin{bmatrix} 1 & 1 \\ 1 & 7 \\ 7 & 20 \end{bmatrix}; \Lambda_B[:, 2] = \begin{bmatrix} 20 & 20 \\ 7 & 1 \\ 1 & 7 \end{bmatrix}$$

- Scenario C - Hard: $\lambda = \Lambda_C$

$$\Lambda_C[:, 1] = \begin{bmatrix} 70 & 12 & 1 \\ 35 & 1 & 35 \\ 1 & 70 & 12 \\ 12 & 35 & 70 \end{bmatrix}; \Lambda_C[:, 2] = \begin{bmatrix} 35 & 70 & 12 \\ 70 & 70 & 70 \\ 12 & 1 & 35 \\ 1 & 70 & 1 \end{bmatrix}; \Lambda_C[:, 3] = \begin{bmatrix} 12 & 70 & 35 \\ 35 & 12 & 70 \\ 70 & 35 & 12 \\ 12 & 1 & 35 \end{bmatrix}$$

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