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A generalized model interpolating between the random energy model and the branching random walk

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Abstract

We study a generalization of the model introduced in [22] that interpolates between the random energy model (REM) and the branching random walk (BRW). More precisely, we are interested in the asymptotic behaviour of the extremal process associated to this model. In [22], Kistler and Schmidt show that the extremal process of the $GREM(N^\alpha)$, $\alpha \in [0, 1)$ converges weakly to a simple Poisson point process. This contrasts with the extremal process of the branching random walk ($\alpha = 1$) which was shown to converge toward a *decorated* Poisson point process by Madaule [20]. In this paper we propose a generalized model of the $GREM(N^\alpha)$, that has the structure of a tree with k_n levels, where $(k_n \leq n)$ is a non-decreasing sequence of positive integers. We show that as long as $\frac{k_n}{n} \rightarrow_{n \rightarrow \infty} 0$, the decoration disappears and we have convergence to a simple Poisson point process. We study a generalized case, where the position of the particles are not necessarily Gaussian variables and the reproduction law is not necessarily binary.

Keywords: Extremal processes, Branching random walk, extremes of log-correlated random fields.

MSC 2020: Primary: 60G80, 60G70, 60G55. Secondary: 60G50, 60G15, 60F05.

1 Introduction

The random energy model (REM) was introduced by Derrida in 1981 [11] for the study of spin glasses. In the REM, there are 2^N spin configurations. Each configuration $\sigma \in \{-1, 1\}^N$ corresponds to an independent centred Gaussian random variable X_σ with variance N , that models its energy level. It is well-known that the extremal process of the REM, which is defined as

$$\mathcal{E}_N = \sum_{\sigma \in \{-1, 1\}^N} \delta_{X_\sigma - m_N}, \quad \text{where } m_N = \beta_c N - \frac{1}{2\beta_c} \log(N) \text{ and } \beta_c = \sqrt{2 \log(2)}, \quad (1)$$

converges weakly in distribution to a Poisson point process with intensity $\frac{1}{\sqrt{2\pi}} e^{-\beta_c x} dx$. Additionally the law of the maximum $M_N = \max_{\sigma \in \{-1, 1\}^N} X_\sigma$ centred by m_N converges weakly to a Gumbel random variable.

Derrida introduced a generalized model in 1985, called the GREM [12], that has the structure of a tree with K levels and can be described as follows. Start by an unique individual (the root). It gives birth to $2^{\frac{N}{K}}$ (we assume that $\frac{N}{K}$ is a positive integer) children at the first level. At each level i , $1 \leq i < K$, each child gives birth independently to $2^{\frac{N}{K}}$ children. We associate each branch of this tree to an independent centred Gaussian random variable with variance $\frac{N}{K}$. In the context of spin glasses, we obtain 2^N configurations in the level K , and the level energy of each configuration is the sum of the values along the branches that forms the path from the root of the tree to the leaf corresponding to this configuration. We call this model $GREM_N(K)$.

Note that the REM can then be thought of as a GREM with one level, i.e. a $GREM_N(1)$. The correlation of the energy of two different configurations depends on the number of common branches shared by their paths from the root up to the node at which they split. These correlations do not have any impact on the extreme values of the energy levels, as the result described in (1) still holds even if $(X_\sigma, \sigma \in \{-1, 1\}^N)$ is distributed as a $GREM_N(K)$, as $N \rightarrow \infty$.

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Kistler and Schmidt [22] studied the asymptotic of the extremal process of a GREM with a number of levels $K_N = N^\alpha$, for $\alpha \in [0, 1)$. They proved that, setting

$$m_N^{(\alpha)} = \beta_c N - \frac{2\alpha + 1}{2\beta_c} \log(N),$$

the extremal process of the $GREM_N(N^\alpha)$ converges weakly to a Poisson point process with intensity $\frac{1}{\sqrt{2\pi}} e^{-\beta_c x} dx$, and the law of the maximum converges to a Gumbel distribution. In the $GREM_N(N^\alpha)$ the stronger correlations between the leaves of the tree have the effect of decreasing the median of the maximal energy level, specifically its logarithmic correction. However the limiting law of the extremal process remains unchanged. In the case of $\alpha = 1$, which corresponds to the classical binary branching random walk, the asymptotic behaviour of the extremal process is well-known. The convergence in law of the recentred maximum was proved by Aidékon [2], and recently Madaule [20] showed the convergence of the extremal process to a decorated Poisson point process with random intensity. Therefore a phase transition can be exhibited, from a simple Poisson point process appearing in the $GREM_N(N^\alpha)$ for $\alpha < 1$ to a decorated one for $\alpha = 1$.

The aim of this article is to have a closer look at this phase transition. We take interest in a generalized version of the $GREM_N(N^\alpha)$, that has the structure of a tree with k_n levels, where $(k_n \geq 0)$ is a non-decreasing sequence of positive integers. We study the asymptotic behaviour of the extremal point process showing that as long as $\frac{k_n}{n} \rightarrow_{n \rightarrow \infty} 0$, the decoration does not appear.

2 Notation and main result

A branching random walk on \mathbb{R} is a particle system that evolves as follows. It starts with a unique individual located at the origin at time 0. At each time $n \geq 1$, each individual alive in the process dies and gives birth to a random number of children, that are positioned around their parent according to i.i.d random variables.

The process we take interest in can be described as follows. Let k_n be an integer sequence growing to ∞ such that $k_n \leq n$ for all $n \in \mathbb{N}$ and set $b_n = \lfloor \frac{n}{k_n} \rfloor$ the integer part of $\frac{n}{k_n}$. The process starts with an unique individual located at the origin at time 0. The particles reproduce for b_n consecutive steps consecutively, each particle giving birth to an i.i.d. number of children. Then each descendant of the initial ancestors moves independently, making b_n i.i.d. steps of displacements. This forms the first generation of the process. For each $1 \leq k \leq k_n$, every individual at generation k repeats independently of the others the same reproduction and displacement procedure as the original ancestor. In other words every individual creates a number of descendants given by the value at time b_n of a Galton-Watson process, whose positions are given by i.i.d. random variables with the same law as a random walk of length b_n .

To describe the model formally we introduce Ulam-Harris notation for trees. Set

$$\mathcal{U} = \bigcup_{n \geq 0} \mathbb{N}^n$$

with $\mathbb{N}^0 = \{\emptyset\}$ by convention. The element (u_1, u_2, \dots, u_n) represents the u_n^{th} child of u_{n-1}^{th} child ..., of u_1 of the root particle which is noted \emptyset . If $u = (u_1, u_2, \dots, u_n)$ we denote by $u_k = (u_1, u_2, \dots, u_k)$ the sequence consisting of the k^{th} first values of u and by $|u|$ the generation of u . For $u, v \in \mathcal{U}$ we denote by $\pi(u)$ the parent of u . If $u = (u_1, u_2, \dots, u_n)$ and $v = (v_1, v_2, \dots, v_n)$, then we write $u.v = (u_1, u_2, \dots, u_n, v_1, v_2, \dots, v_n)$ for the concatenation of u and v . We write

$$|u \wedge v| := \inf\{j \leq n : u_j = v_j \text{ and } u_{j+1} \neq v_{j+1}\}.$$

This quantity is called the overlap of u and v in the context of spin glasses. A tree \mathcal{T} is a subset of \mathcal{U} satisfying the following assumptions:

- $\emptyset \in \mathcal{T}$.
- if $u \in \mathcal{T}$, then $\pi(u) \in \mathcal{T}$.
- if $u = (u_1, u_2, \dots, u_n) \in \mathcal{T}$, then $\forall j \leq n, \pi(u).j \in \mathcal{T}$.

We now introduce the reproduction and displacement laws associated to our process. Let $(Y_n)_{n \in \mathbb{N}}$ be a random walk such that $\mathbb{E}(Y_1) = 0$ and $\text{Var}(Y_1) = 1$. We denote by $(Z_n)_{n \in \mathbb{N}}$ a Galton-Watson process such that $Z_0 = 1$ and offspring law given by the weights $(p(k))_{k \in \mathbb{N}}$ with $p_0 = 0$. Under this assumption, the Galton Watson process survives almost surely. Set $m = \sum_{k \geq 1} kp(k)$ the mean of the offspring distribution and assume that $m > 1$. Recall that the Galton-Watson process $(Z_n)_{n \in \mathbb{N}}$ satisfies for all $n \in \mathbb{N}$:

$$Z_{n+1} = \sum_{j=1}^{Z_n} \xi_{n+1,j},$$

where $(\xi_{n,j})_{1 \leq j \leq Z_n}$ are i.i.d random variables with law $(p(k))_{k \in \mathbb{N}}$.

Under the assumption $\mathbb{E}(Z_1 \log(Z_1)) < \infty$, Kesten and Stigum [17] proved that on the set of non extinction of \mathcal{T} there exists a positive random variable Z_∞ such that

$$\lim_{b \rightarrow \infty} \frac{Z_b}{m^b} = Z_\infty > 0, \quad \text{a.s.} \quad (2)$$

In this article we assume that the following stronger condition holds:

$$\mathbb{E}(Z_1^2) < \infty. \quad (3)$$

Construct a tree that we denote $\mathcal{T}^{(n)}$ as follows. Start by the ancestor \emptyset located at the origin. It gives birth to Z_{b_n} children. For each $k \leq k_n$, each individual at the generation k gives birth to an independent copy of Z_{b_n} , that are positioned according to i.i.d random variables with the same law as Y_{b_n} . For $1 \leq k \leq k_n$, let

$$\mathcal{H}_k := \{u \in \mathcal{T}^{(n)} : |u| = k\},$$

the set of particles in the k^{th} generation. By construction, we have $\#\mathcal{H}_k = Z_{kb_n}$ in law for all $k \leq k_n$. We define $(X_u^{(n)}, u \in \mathcal{T}^{(n)})$ a family of i.i.d. random variables with same law as Y_{b_n} . For $u \in \mathcal{T}^{(n)}$, we write

$$S_u^{(n)} = \sum_{k=1}^{|u|} X_{u_k}^{(n)}.$$

The goal of this paper is to study the asymptotic behaviour of the extremal process associated to this model

$$\mathcal{E}_n^{(b_n)} = \sum_{u \in \mathcal{H}_{k_n}} \delta_{S_u^{(n)} - m_n}.$$

Let us introduce notation associated to the displacement of the process. For all $\theta > 0$ we set

$$\Lambda(\theta) := \log(\mathbb{E}(\exp(\theta Y_1))). \quad (4)$$

We assume that there exists $\theta > 0$ such that $\Lambda(\theta) < \infty$. We write:

$$\kappa_n(\theta) = \log \mathbb{E} \left(\sum_{|u|=1} e^{\theta X_u^{(n)}} \right).$$

Observe that $\kappa_n(\theta) = b_n(\log(m) + \Lambda(\theta))$ as

$$\mathbb{E} \left(\sum_{|u|=1} e^{\theta X_u^{(n)}} \right) = \mathbb{E} \left(\sum_{|u|=1} \mathbb{E}(e^{\theta X_u^{(n)}} | Z_{b_n}) \right) = \mathbb{E}(Z_{b_n} \mathbb{E}(e^{\theta Y_{b_n}})) = e^{b_n(\log(m) + \Lambda(\theta))}.$$

The function κ_n is convex and differentiable on $\{\theta > 0, \kappa_n(\theta) < \infty\}$, its interval of definition. We assume that there exists $\theta^* > 0$ such that

$$\theta^* \Lambda'(\theta^*) - \Lambda(\theta^*) = \log(m). \quad (5)$$

We also assume that there exists $\delta > 0$ such that

$$\mathbb{E}(\exp((\theta^* + \delta)Y_1)) < \infty \quad (6)$$

Recall that the case $k_n = n$ corresponds to the classical branching random walk. Then under assumption (4) and (5), Kingman [18], Hammersley [14] and Biggins [7] showed that on the set of non-extinction of \mathcal{T}

$$\lim_{n \rightarrow \infty} \frac{M_n}{n} := \frac{\kappa(\theta^*)}{\theta^*} = v \quad \text{a.s.},$$

where, $M_n = \max_{u \in \mathcal{H}_n} S_u$ and v is the speed of the right-most individual. Then, Hu and Shi [15] and Addario-Berry and Reed [1] proved that

$$M_n = nv - \frac{3}{2\theta^*} \ln(n) + O_{\mathbb{P}}(1),$$

where $O_{\mathbb{P}}(1)$ represents a tight sequence of random variables.

Throughout this paper we will assume that we are in one of the two cases:

(**H**₁): Y_1 is a standard Gaussian variable and $b_n \rightarrow \infty$ as $n \rightarrow \infty$.

(**H**₂): The characteristic function $\phi(\lambda) = \mathbb{E}(\exp(i\lambda Y_1))$ of Y_1 satisfies the Cramér condition, i.e

$$\limsup_{|\lambda| \rightarrow \infty} |\phi(\lambda)| < 1,$$

and $\frac{b_n}{\log(n)^2} \rightarrow_{n \rightarrow \infty} \infty$ as $n \rightarrow \infty$.

Our work is inspired by the recent works on the convergence of the extremal processes [4], [5], [22] and [20]. The main result of this paper is the following convergence in distribution.

Theorem 1. *Assume that (3), (4), (5), (6) and either (**H**₁) or (**H**₂) hold, then setting*

$$m_n = k_n b_n v - \frac{3}{2\theta^*} \log(n) + \frac{\log(b_n)}{\theta^*},$$

the extremal process

$$\mathcal{E}_n^{(b_n)} = \sum_{u \in \mathcal{H}_{k_n}} \delta_{S_u^{(n)} - m_n}$$

converges in law to a Poisson point process with intensity $\frac{1}{\sqrt{2\pi\sigma^2}} Z_\infty e^{-\theta^* x}$, where $\sigma^2 = \kappa_n''(\theta^*)$ and Z_∞ is the random variable defined in equation (2). Moreover, the law of the recentered maximum converges weakly to a randomly Gumbel distribution shifted by $\frac{1}{\theta^*} \log(Z_\infty)$.

Remark 2. Denote by $\mathcal{C}_b^{l,+}$ the set of continuous, positive and bounded functions $\phi : \mathbb{R} \rightarrow \mathbb{R}_+$ with support bounded on the left. By [6, Lemma 4.1], it is enough to show that for all function $\phi \in \mathcal{C}_b^{l,+}$

$$\lim_{n \rightarrow \infty} \mathbb{E} \left(e^{-\sum_{u \in \mathcal{H}_{k_n}} \phi(S_u^{(n)} - m_n)} \right) = \mathbb{E} \left(\exp \left(-Z_\infty \frac{1}{\sqrt{2\pi\sigma^2}} \int e^{-\theta^* y} (1 - e^{-\phi(y)}) dy \right) \right).$$

The result of Kistler and Schmidt [22, Theorem 1.1] is covered by Theorem 2. It is the case (**H**₁) with $k_n = N^\alpha$, $0 \leq \alpha < 1$ and $Z_1 = 2$ in our theorem. In that case we have $Z_\infty = 1$ and $m_n = n\beta_c - \frac{2\alpha+1}{2\beta_c} \log(n)$. Throughout this paper, we use C and c to denote a generic positive constants, that may change from line to line. We say that $f_n \sim_{n \rightarrow \infty} g_n$ if $\lim_{n \rightarrow \infty} \frac{f_n}{g_n} = 1$. For $x \in \mathbb{R}$ we write $x_+ = \max(x, 0)$.

The rest of the paper is organized as follows. In the next section, we introduce the many to one lemma, and we will give a series of useful random walk estimates. In Section 4 we introduce a modified extremal process which we show to have same asymptotic behaviour of the original extremal process defined in the principal theorem. Finally we will conclude the paper with a proof of the main result.

3 Many-to-one formula and random walk estimates

In this section, we introduce the many-to-one lemma, that links additive moments of branching processes to random walk estimates. We then introduce some estimates for the asymptotic behaviour of random walks conditioned to stay below a line, and prove their extension to a generalized random walk where the law of each step is given by the sum of b_n i.i.d random variables.

3.1 Many-to-one formula

We start by introducing the celebrated many-to-one lemma that transforms an additive function of a branching random walk into a simple function of random walk. This lemma was introduced by Kahane and Peyrière [16]. Before we introduce it, we need to define some change of probabilities and to introduce some notation.

Let $W_0 := 0$ and $(W_j - W_{j-1})_{j \geq 1}$ be a sequence of independent and identically distributed random variables such that for any measurable function $h : \mathbb{R} \mapsto \mathbb{R}$,

$$\mathbb{E}(h(W_1)) = \mathbb{E} \left(e^{\theta^* Y_1 - \Lambda(\theta^*)} h(Y_1) \right).$$

where Y_1 is the law defined in Section 2. Respectively, we introduce $(T_j^{(n)} - T_{j-1}^{(n)})_{j \geq 1}$ a sequence of i.i.d random variables such that $T_0 = 0$ and

$$\mathbb{E}(h(T_1^{(n)})) = \frac{\mathbb{E} \left(\sum_{u, |u|=1} e^{\theta^* S_u^{(n)}} h(S_u^{(n)}) \right)}{\mathbb{E} \left(\sum_{u, |u|=1} e^{\theta^* S_u^{(n)}} \right)} = \mathbb{E} \left(e^{\theta^* Y_{b_n} - \Lambda(\theta^*)} h(Y_{b_n}) \right). \quad (7)$$

Observe that $(T_k^{(n)}, k \geq 1)$ is a sequence of random variables that have the same law as the process $(U_{kb_n} = \sum_{j=1}^{kb_n} W_j, k \geq 1)$. We now set $\bar{T}_j^{(n)} = T_j^{(n)} - j b_n v$ respectively $\bar{W}_j = W_j - j v, j \geq 1$. We have

$$\mathbb{E}(W_1) = \mathbb{E} \left(Y_1 e^{\theta^* Y_1 - \Lambda(\theta^*)} \right) = \Lambda'(\theta^*),$$

and as $\Lambda'(\theta^*) = \kappa_n'(\theta^*) = v$, we have $\mathbb{E}(\bar{W}_1) = 0$ and similarly

$$\mathbb{E}(W_1^2) = \mathbb{E} \left(Y_1^2 e^{\theta^* Y_1 - \Lambda(\theta^*)} \right) = \Lambda''(\theta^*) + (\Lambda'(\theta^*))^2,$$

which gives $\text{Var}(\bar{W}_1) = \Lambda''(\theta^*) = \sigma^2$ which is finite by assumption (6). As a consequence we have $\mathbb{E}(\bar{T}_1^{(n)}) = 0$ and $\text{Var}(\bar{T}_1^{(n)}) = b_n \sigma^2 < \infty$. In the case (\mathbf{H}_1) , note that \bar{W}_1 is a standard Gaussian random variable which mean that $\bar{T}_1^{(n)}$ is a centred Gaussian random variable with variance b_n .

For simplicity we write S_u in place of $S_u^{(n)}$ and T_j in place of $T_j^{(n)}$ in the rest of the article.

Proposition 3. [23, Theorem 1.1] *For any $j \geq 1$ and any measurable function $g : \mathbb{R}^j \mapsto \mathbb{R}_+$, we have*

$$\mathbb{E} \left(\sum_{|u|=j} g((S_{u_i})_{1 \leq i \leq j}) \right) = \mathbb{E} \left(e^{-\theta^* \bar{T}_i} g((\bar{T}_i + i b_n v)_{1 \leq i \leq j}) \right).$$

Proof. For $j = 1$, by (7) and using that $b_n v = \frac{\kappa_n(\theta^*)}{\theta^*}$, we have

$$\mathbb{E} \left(\sum_{|u|=1} g(S_u) \right) = \mathbb{E} \left(e^{-\theta^* T_1 + \kappa_n(\theta^*)} g(T_1) \right) = \mathbb{E} \left(e^{-\theta^* \bar{T}_1} g(\bar{T}_1 + b_n v) \right)$$

where $\bar{T}_1 = T_1 - b_n v$. We complete the proof by induction in the the same way as in [23, Theorem 1.1]. \square

3.2 Random walk estimates

In this section we introduce some estimates for the asymptotic behaviour of functionals of the random walks, such as the probability to stay above a boundary. We first give an estimate for the probability that a random walk stays above a boundary $(f_n)_{n \in \mathbb{N}}$, that is $O(n^{1/2-\epsilon})$ for some $\epsilon > 0$. This lemma was introduced in [21, Lemma 3.2].

Lemma 4. *Let $(w_n)_{n \in \mathbb{N}}$ be a centred random walk with finite variance. Fix $\epsilon > 0$, there exists $C > 0$ such that*

$$\mathbb{P}(w_k \geq -(k^{1/2-\epsilon} - y), k \leq n) \leq C \frac{1-y}{\sqrt{n}}$$

for any $y > 0$.

From now on we use the random walks $(T_k)_{k \geq 1}$ and $(\bar{T}_k)_{k \geq 1}$ defined in (7), unless otherwise stated. We introduce a version of the Stone's local limit theorem [25] that gives an approximation of the probability for a random walk to end up in a finite interval.

Lemma 5. *Let $f \in C_b^{l,+}$ be a Riemann integrable function, and let $(r_n)_{n \in \mathbb{N}}$ be a sequence of positive real numbers, such that $\lim_{n \rightarrow \infty} \frac{r_n}{\sqrt{n}} = 0$. Set*

$$a_n = \frac{-3}{2\theta^*} \log(n) + \frac{\log(b_n)}{\theta^*}$$

then we get

$$\mathbb{E}(f(\bar{T}_{k_n} - a_n + x)e^{-\theta^* \bar{T}_{k_n}}) = \frac{e^{\theta^* x n^{3/2}}}{b_n \sqrt{2\pi\sigma^2 k_n b_n}} \int f(y) e^{-\theta^* y} dy (1 + o(1))$$

uniformly in $x \in [-r_n, r_n]$.

Proof. By setting $h(z) = e^{-\theta^* z} f(z)$, it is enough to prove that

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x)) = \frac{1}{\sqrt{2\pi\sigma^2 k_n b_n}} \int h(y) dy (1 + o(1)) \quad (8)$$

uniformly in $x \in [-r_n, r_n]$. We prove this lemma by successive approximations of the function h , starting with an indicator function. Set $h(z) = \mathbf{1}_{[a,b]}(z)$ for some $a < b \in \mathbb{R}$, then we write

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x)) = \mathbb{P}(\bar{T}_{k_n} - a_n + x \in [a, b]), \quad (9)$$

As \bar{T}_1 is the sum of b_n i.i.d. copies of \bar{Z}_1 , \bar{T}_{k_n} is the sum of $k_n b_n$ i.i.d. centred random variables with finite variance, therefore we can apply the Stone's local limit theorem [25] to obtain

$$\mathbb{P}(\bar{T}_{k_n} - a_n + x \in [a, b]) = \frac{b-a}{\sqrt{2\pi\sigma^2 k_n b_n}} \exp\left(\frac{-(a_n - x)^2}{2k_n b_n \sigma^2}\right) (1 + o(1)) = \frac{b-a}{\sqrt{2\pi k_n b_n \sigma^2}} (1 + o(1)),$$

uniformly in $x \in [-r_n, r_n]$, which completes the proof of (8) in that case.

We now assume that h is a continuous function with compact support, we prove (8) by approximating it by scale functions. Denote by $[a, b]$ the support of h . Let $(t_i)_{0 \leq i \leq m}$ be an uniform subdivision of $[a, b]$ where $m \in \mathbb{N}$ is the number of the subdivisions and $t_i = a + i(b-a)/m$ for $0 \leq i \leq m$. Set

$$\underline{h}_m(x) = \sum_{i=0}^{m-1} m_i \mathbf{1}_{\{x \in [t_i, t_{i+1}]\}} \quad \text{and} \quad \bar{h}_m(x) = \sum_{i=0}^{m-1} M_i \mathbf{1}_{\{x \in [t_i, t_{i+1}]\}},$$

where $M_i = \sup_{z \in [t_i, t_{i+1}]} h(z)$ and $m_i = \inf_{z \in [t_i, t_{i+1}]} h(z)$. Hence using the Riemann sum approximation and the fact that f is a non-negative function, for all $\epsilon > 0$, there exists m_0 such that for all $m \geq m_0$ we have

$$(1 - \epsilon) \int_a^b h(y) dy \leq \int_a^b \underline{h}_m(y) dy \leq \int_a^b \bar{h}_m(y) dy \leq (1 + \epsilon) \int_a^b h(y) dy, \quad (10)$$

where $\int_a^b \underline{h}_m(y) dy = \sum_{i=0}^{m-1} \frac{b-a}{m} m_i$ and $\int_a^b \bar{h}_m(y) dy = \sum_{i=0}^{m-1} \frac{b-a}{m} M_i$.

Using equation (9) we have

$$\begin{aligned} \mathbb{E}(\bar{h}_m(\bar{T}_{k_n} - a_n + x)) &= \sum_{i=0}^{m-1} M_i \mathbb{P}(\bar{T}_{k_n} - a_n + x \in [t_i, t_{i+1}]) = \frac{1}{\sqrt{2\pi\sigma^2 k_n b_n}} \sum_{i=0}^{m-1} \frac{b-a}{m} M_i (1 + o(1)) \\ &= \frac{1}{\sqrt{2\pi\sigma^2 k_n b_n}} \int_a^b \bar{h}_m(y) dy (1 + o(1)). \end{aligned}$$

Therefore, using that $\mathbb{E}(h(\bar{T}_k - a_n + x)) \leq \mathbb{E}(\bar{h}_m(\bar{T}_k - a_n + x))$ and by (10) we deduce that

$$\limsup_{n \rightarrow \infty} \sup_{x \in [0, r_n]} \sqrt{k_n b_n} \mathbb{E}(h(\bar{T}_{k_n} - a_n + x)) \leq (1 + \epsilon) \frac{1}{\sqrt{2\pi\sigma^2}} \int_a^b h(y) dy.$$

Using similar arguments we have

$$\liminf_{n \rightarrow \infty} \inf_{x \in [0, r_n]} \sqrt{k_n b_n} \mathbb{E} (h(\bar{T}_{k_n} - a_n + x)) \geq (1 - \epsilon) \frac{1}{\sqrt{2\pi\sigma^2}} \int_a^b h(y) dy.$$

Finally, letting $\epsilon \rightarrow 0$ completes the proof of (8) when h is a compactly support function. Finally we consider the general case, and assume that f is bounded with bounded support on the left. We introduce the function

$$\chi(u) = \begin{cases} 1 & \text{if } u < 0 \\ 1 - u & \text{si } 0 \leq u \leq 1 \\ 0 & \text{if } u > 1 \end{cases}$$

then we write,

$$\begin{aligned} \mathbb{E}(h(\bar{T}_{k_n} - a_n + x)) &= \mathbb{E} (h(\bar{T}_{k_n} - a_n + x)\chi(\bar{T}_{k_n} - a_n + x - B)) \\ &+ \mathbb{E} (h(\bar{T}_{k_n} - a_n + x)(1 - \chi(\bar{T}_{k_n} - a_n + x - B))) \end{aligned}$$

for some $B > 0$. Observe that the function $z \mapsto h(z)\chi(z - B)$ is continuous with compact support as a consequence we have

$$\begin{aligned} \mathbb{E}(h(\bar{T}_{k_n} - a_n + x)) &= \frac{1}{\sqrt{2\pi\sigma^2 k_n b_n}} \int h(y)\chi(y - B) dy (1 + o(1)) \\ &+ \mathbb{E} (h(\bar{T}_{k_n} - a_n + x)(1 - \chi(\bar{T}_{k_n} - a_n + x - B))). \end{aligned} \quad (11)$$

Using the Stone's local limit theorem [25] there exists a constant $C > 0$ such that the second quantity in the right-hand side of (11) is bounded by

$$\begin{aligned} \mathbb{E} (h(\bar{T}_{k_n} - a_n + x)(1 - \chi(\bar{T}_{k_n} - a_n + x - B))) &\leq \mathbb{E} \left(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_{k_n} - a_n + x > B\}} \right) \\ &\leq \|f\|_\infty \mathbb{E} \left(\sum_{j \geq B} e^{-\theta^* j} \mathbf{1}_{\{\bar{T}_{k_n} - a_n + x \in [j, j+1]\}} \right) \leq C \|f\|_\infty \frac{e^{-\theta^* B}}{\sqrt{k_n b_n \sigma^2}}. \end{aligned}$$

On the other hand by the dominated convergence theorem we have

$$\lim_{B \rightarrow \infty} \frac{1}{\sqrt{2\pi k_n b_n}} \int h(y)\chi(y - B) dy = \frac{1}{\sqrt{2\pi\sigma^2 k_n b_n}} \int_0^\infty h(y) dy,$$

Now using similar arguments to those used in the last case we deduce that

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x)) = \frac{1}{\sqrt{2\pi\sigma^2 k_n b_n}} \int f(y) e^{-\theta^* y} dy (1 + o(1)),$$

which completes the proof. \square

3.2.1 Random walk with Gaussian steps

In this section we assume that (\mathbf{H}_1) holds, i.e that $(\bar{T}_k)_{k \geq 0}$ is a Gaussian random walk. Let $(\beta_n(k), k \leq k_n)$ be the standard discrete Brownian bridge with k_n steps, which can be defined as,

$$\beta_n(k) = \frac{1}{\sqrt{b_n}} \left(\bar{T}_k - \frac{k}{k_n} \bar{T}_{k_n} \right).$$

In the following lemma we estimate the probability for a Brownian bridge to stay below a boundary during all his lifespan. This lemma was introduced in [9].

Lemma 6. *Let h be the function defined by*

$$h(k) = \begin{cases} 0 & \text{if } k = 0 \text{ or } k = k_n \\ a \log((k_n - k) \wedge k) b_n + 1 & \text{otherwise.} \end{cases}$$

where a is a positive constant. There exists a constant $C > 0$ such that for all $x > 0$ and $n \geq 0$ we have

$$\mathbb{P} \left(\beta_n(k) \leq \frac{1}{\sqrt{b_n}} (h(k) + x), k \leq k_n \right) \leq C \frac{(1 + \frac{x}{\sqrt{b_n}})^2}{k_n}. \quad (12)$$

We refer to the function $k \mapsto h(k)$ as a barrier. An application of this lemma is to give an upper bound for the probability that a random walk with Gaussian steps make an excursion above a well-chosen barrier.

Lemma 7. *Let $\alpha > 0$, and for $0 \leq k \leq k_n$ we write $f_n(k) = \alpha \log(\frac{(k_n-k)b_n+1}{k_nb_n})$. There exists $C > 0$ such that for all $x \geq 0$, $a < b \in \mathbb{R}$ and $k \leq k_n$ we have*

$$\mathbb{P}(\bar{T}_k - f_n(k) \in [a, b], \bar{T}_j \leq f_n(j) + x, j \leq k) \leq C(b-a) \frac{(1 + \frac{x}{\sqrt{b_n}})^2}{\sqrt{b_n} k^{\frac{3}{2}}}.$$

Proof. For $n \in \mathbb{N}$ we have

$$\begin{aligned} & \mathbb{P}(\bar{T}_k - f_n(k) \in [a, b], \bar{T}_j \leq f_n(j) + x, j \leq k) \\ & \leq \mathbb{P}\left(\bar{T}_k - f_n(k) \in [a, b], \bar{T}_j - \frac{j}{k}\bar{T}_k \leq f_n(j) + x - \frac{j}{k}(f_n(k) + a), j \leq k\right), \end{aligned}$$

using independence between the discrete Brownian bridge $\bar{T}_j - \frac{j}{k}\bar{T}_k$ and \bar{T}_k we obtain

$$\begin{aligned} & \mathbb{P}\left(\bar{T}_k - f_n(k) \in [a, b], \bar{T}_j - \frac{j}{k}\bar{T}_k \leq f_n(j) + x - \frac{j}{k}(f_n(k)), j \leq k\right) \\ & \leq \mathbb{P}(\bar{T}_k - f_n(k) \in [a, b]) \mathbb{P}\left(\bar{T}_j - \frac{j}{k}\bar{T}_k \leq f_n(j) + x - \frac{j}{k}(f_n(k) + a), j \leq k\right). \end{aligned} \quad (13)$$

To estimate the probability that a discrete Brownian bridge stay below a logarithmic barrier, we apply Lemma 6. First observe that the function $x \mapsto \frac{\log(x)}{x}$ is decreasing for $x \geq e$, and using that $k_n - j + 1 \leq (k_n - k + 1) + (k - j) + 1 \leq 2(k_n - k + 1)(k - j + 1)$, we have for $j \leq \frac{k}{2}$,

$$\begin{aligned} f_n(j) + x - \frac{j}{k}(f_n(k) + a) & \leq \alpha \frac{j}{k} \left(\log\left(\frac{k_nb_n}{(k_n-k)b_n+1}\right) - \log\left(\frac{k_nb_n}{(k_n-j)b_n+1}\right) \right) + x \\ & \leq \alpha \frac{j}{k} (\log(kb_n) + \log(2)) + x \leq \alpha (\log((jb_n \vee e)) + \log(2)) + x \end{aligned}$$

and for $\frac{k}{2} \leq j \leq k$, we have

$$\begin{aligned} f_n(j) + x - \frac{j}{k}(f_n(k) + a) & \leq \alpha (\log(\frac{k_nb_n}{(k_n-k)b_n+1}) + x - \log(\frac{k_nb_n}{(k_n-j)b_n+1})) \\ & \leq \alpha (\log(((k_n-j)b_n+1) - \log((k_n-k)b_n+1))) + x \\ & \leq \alpha (\log(2) + \log(1 + (k-j)b_n)) + x. \end{aligned}$$

Then by Lemma 6 we get after rescaling by $\frac{1}{\sqrt{b_n}}$ the following upper bound

$$\begin{aligned} & \mathbb{P}\left(\bar{T}_j - \frac{j}{k}\bar{T}_k \leq f_n(k) - \frac{j}{k}(f_n(j) - x), j \leq k\right) \\ & \leq \mathbb{P}\left(\beta_n(k) \leq \alpha (\log((k \wedge (k-j)) + 1)) + \frac{x}{\sqrt{b_n}} + 1, j \leq k\right) \leq C \frac{(1 + \frac{x}{\sqrt{b_n}})^2}{k}, \end{aligned}$$

where C is a positive constant. To bound the first quantity in (13) we use the Gaussian estimate

$$\mathbb{P}(\bar{T}_k - f_n(k) \in [a, b]) \leq \frac{b-a}{\sqrt{kb_n}}$$

which completes the proof. \square

From now we denote by $B_n(k) = \frac{\bar{T}_k}{\sqrt{b_n}}$. Recall that under (\mathbf{H}_1) , $(B_n(k))_{k \leq k_n}$ is a standard random walk with i.i.d Gaussian steps. Define the function $L : (0, \infty) \mapsto (0, \infty)$ by $L(0) = 1$ and

$$L(x) := \sum_{k \geq 0} \mathbf{P}\left(B_n(k) \geq -x, B_n(k) \leq \min_{j \leq k-1} (B_n(j))\right) \quad \text{for } x > 0.$$

It is known by [13, section XII.7], that the function L is the renewal function associated to the random walk $(B_n(k))_{k \geq 0}$. We will cite some properties that are mentioned in [13, section XII.7]. The fundamental property of the renewal function is

$$L(x) = \mathbb{E} \left(L(x + B_n(1)) \mathbf{1}_{\{x + B_n(1) \geq 0\}} \right), \quad (14)$$

and is a right-continuous and non-decreasing function. Since in case (\mathbf{H}_1) , the initial law has no atoms, then the function L is continuous. Also, there exists a constant $c_0 > 0$ such that

$$\lim_{x \rightarrow \infty} \frac{L(x)}{x} = c_0. \quad (15)$$

In particular there exists a constant $C > 0$ such that for all $x \in \mathbb{R}$

$$L(x) \leq C(1 + x_+). \quad (16)$$

Also we have by, for $x, y \geq 0$

$$L(x + y) \leq 2L(x)L(y). \quad (17)$$

Similarly, we define $L_-(x)$ as the renewal function associated to $-B$.

Since \bar{T} is a symmetric law we have $L_-(x) = L(x)$ for all $x \geq 0$. It is also known that there exists a positive constant C_1 such that for $y \geq 0$

$$\mathbb{P} \left(\min_{k \leq k_n} (B_n(k)) \geq -y \right) \sim_{n \rightarrow \infty} C_1 \frac{L(y)}{\sqrt{k_n}}. \quad (18)$$

By Theorem 3.5 in [24], assuming that B is Gaussian we have $C_1 = \frac{1}{\sqrt{\pi}}$. We now introduce an approximation of the probability for a random walk to stay below a line and end up in a finite interval. Set

$$\tilde{F}_n(k) = \frac{k}{k_n} a_n = \frac{k}{k_n} (m_n - k_n b_n v), \quad k = 0, \dots, k_n, \quad n \in \mathbb{N}.$$

Lemma 8. *Let $(r_n)_{n \in \mathbb{N}}$ be a sequence of positive real numbers such that $\lim_{n \rightarrow \infty} \frac{r_n}{\sqrt{k_n}} = 0$. Let*

$$a_n = \frac{-3}{2\theta^*} \log(n) + \frac{\log(b_n)}{\theta^*}.$$

For all $f \in \mathcal{C}_b^{l,+}$ we have

$$\mathbb{E} \left(f(\bar{T}_{k_n} - a_n + x) e^{-\theta^* \bar{T}_{k_n}} \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}} \right) = \frac{e^{\theta^* x}}{\sqrt{2\pi}} \int_{-\infty}^0 f(y) e^{-\theta^* y} dy \left(R\left(\frac{-x}{\sqrt{b_n}}\right) + o(1) \right).$$

uniformly in $x \in [-r_n, 0]$.

Proof. By setting $h(z) = e^{-\theta^* z} f(z)$ it is enough to prove that

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}}) = \frac{1}{k_n^{3/2} \sqrt{2\pi b_n}} \int_{-\infty}^0 h(y) dy \left(R\left(\frac{-x}{\sqrt{b_n}}\right) + o(1) \right) \quad (19)$$

uniformly in $x \in [-r_n, 0]$.

Following the same method used in Lemma 5 it is enough to prove this estimate for an indicator function. By writing $\mathbf{1}_{[-a, -b]} = \mathbf{1}_{[-a, 0]} - \mathbf{1}_{[-b, 0]}$ for some $a > 0, b > 0$, it is enough to prove this estimate for $h(z) = \mathbf{1}_{[-a, 0]}(z)$, in that case we have

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}}) = \mathbb{P}(\bar{T}_{k_n} - a_n + x \geq -a, \bar{T}_k \leq F_n(k) - x, k \leq k_n).$$

Define a new probability measure \mathbb{Q} on \mathbb{R} by

$$\frac{d\mathbb{P}}{d\mathbb{Q}}(\bar{T}) = \exp\left(\frac{-a_n}{n} \bar{T} + \Lambda\left(\frac{a_n}{n}\right)\right) \quad (20)$$

where $\Lambda(\theta) = \frac{\theta^2}{2}$. Then we rewrite

$$\begin{aligned} & \mathbb{P}(\bar{T}_{k_n} - a_n + x \geq -a, \bar{T}_k \leq F_n(k) - x, k \leq k_n) \\ &= \mathbb{E}_{\mathbb{Q}}\left(e^{-\frac{a_n}{n}(\sqrt{b_n}\hat{B}_n(k_n) - \frac{a_n}{2n})} \mathbf{1}_{\{\sqrt{b_n}\hat{B}_n(k_n) + x \geq -a, \sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n\}}\right), \end{aligned}$$

where $\hat{B}_n(k) = B_n(k) - \frac{k}{\sqrt{b_n k_n}} a_n$. Observe that the law of \hat{T} under \mathbb{Q} is the same as the law of \bar{T} under \mathbb{P} .

Under this change of measure, we can rewrite the probability as

$$\begin{aligned} & \mathbb{E}_{\mathbb{Q}}\left(e^{-\frac{a_n}{n}\sqrt{b_n}\hat{B}_n(k_n) + \frac{a_n^2}{2n}} \mathbf{1}_{\{\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n, \sqrt{b_n}\hat{B}_n(k_n) + x \geq -a\}}\right) \\ & \leq e^{\frac{a_n}{n}(x+a) + \frac{a_n^2}{2n}} \mathbb{Q}\left(\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n, \sqrt{b_n}\hat{B}_n(k_n) \geq -a - x\right). \end{aligned}$$

as a consequence

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, 0]} \mathbb{E}_{\mathbb{Q}}\left(e^{-\frac{a_n}{n}\sqrt{b_n}\hat{B}_n(k_n) + \frac{a_n^2}{2n}} \mathbf{1}_{\{\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n, \sqrt{b_n}\hat{B}_n(k_n) + x \geq -a\}}\right) \\ & \leq \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, 0]} \mathbb{Q}\left(\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n, \sqrt{b_n}\hat{B}_n(k) \geq -a - x\right), \end{aligned}$$

similarly we have

$$\begin{aligned} & \liminf_{n \rightarrow \infty} \inf_{x \in [-r_n, 0]} \mathbb{E}_{\mathbb{Q}}\left(e^{-\frac{a_n}{n}\sqrt{b_n}\hat{B}_n(k_n) + \frac{a_n^2}{2n}} \mathbf{1}_{\{\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n, \sqrt{b_n}\hat{B}_n(k_n) + x \geq -a\}}\right) \\ & \leq \liminf_{n \rightarrow \infty} \inf_{x \in [-r_n, 0]} \mathbb{Q}\left(\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n, \sqrt{b_n}\hat{B}_n(k_n) \geq -a - x\right). \end{aligned} \quad (21)$$

for all $a > 0$. Therefore, it remains to estimate the quantity (21). Applying the Markov property at time $p = \lfloor \frac{k_n}{2} \rfloor$ we get

$$\mathbb{Q}\left(\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n, \sqrt{b_n}\hat{B}_n(k_n) \geq -a - x\right) = \mathbb{E}\left(f_{x,n,a}(\sqrt{b_n}\hat{B}_n(p)) \mathbf{1}_{\{\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq p\}}\right) \quad (22)$$

where for all $y \leq 0$

$$f_{x,n,a}(y) = \mathbb{Q}\left(\sqrt{b_n}\hat{B}_n(k_n - p) + y \geq -a - x, \sqrt{b_n}\hat{B}_n(k) + y \leq -x, k \leq k_n - p\right).$$

Using that the process $(\sqrt{b_n}(\hat{B}_n(k_n - p) - \hat{B}_n(k_n - p - j)), 0 \leq j \leq k_n - p)$ has the same law as $(\sqrt{b_n}\hat{B}_n(j), 0 \leq j \leq k_n - p)$ under \mathbb{Q} , we obtain

$$\begin{aligned} f_{x,n,a}(y) &= \mathbb{Q}\left((-\sqrt{b_n}\hat{B}_n(k)) \leq (-\sqrt{b_n}\hat{B}_{k_n-p}) - (x+y) \leq a, k \leq k_n - p\right) \\ &= \mathbb{Q}\left(\sqrt{b_n}\hat{B}_n(k) \leq \sqrt{b_n}\hat{B}_n(k_n - p) - (x+y) \leq a, k \leq k_n - p\right) \end{aligned}$$

since $(\sqrt{b_n}\hat{B}_n(k))_{k \geq 0}$ is a symmetric law. We write $\check{B}_n(k_n - p) = \max_{0 \leq j \leq k_n - p} \sqrt{b_n}\hat{B}_n(j)$, set

$$\tau_{k_n - p} = \min\left\{i : 0 \leq i \leq k_n - p, \check{B}_n(k_n - p) = \sqrt{b_n}\hat{B}_n(i)\right\}$$

the first time when $\sqrt{b_n}\hat{B}_n(i)$ hits its maximum in the interval $[0, k_n - p]$. We have

$$f_{x,n,a}(y) = \sum_{i=0}^{k_n - p} \mathbb{Q}\left(\tau_{k_n - p} = i, \sqrt{b_n}\hat{B}_n(k) \leq \sqrt{b_n}\hat{B}_n(k_n - p) - (x+y) \leq a, k \leq k_n - p\right).$$

Applying the Markov property at time i we get

$$f_{x,n,a}(y) = \sum_{i=0}^{k_n - p} \mathbb{E}\left(g(\check{B}_n(i) - a) \mathbf{1}_{\{\check{B}_n(i) = \sqrt{b_n}\hat{B}_n(i) \leq a\}}\right),$$

where for all $z \leq 0$, $g_{x,n,y}(z) = \mathbb{Q}\left(y + x \leq \sqrt{b_n}\hat{B}_n(k_n - p - i) \leq y + x - z, \check{B}_n(k_n - p - i) \leq 0\right)$.

We now split the sum $\sum_{i=0}^{k_n-p}$ into $\sum_{i=0}^{i_n} + \sum_{i=i_n+1}^{k_n-p}$, where $i_n = \lfloor \sqrt{k_n} \rfloor$, then we write

$$f_{n,x,a}(y) = f_{n,x,a}^{(1)}(y) + f_{n,x,a}^{(2)}(y)$$

where

$$f_{n,x,a}^{(1)}(y) = \sum_{i=0}^{i_n} \mathbb{E}\left(g(\check{B}_n(i) - a) \mathbf{1}_{\{\check{B}_n(i) = \sqrt{b_n}\hat{B}_n(i) \leq a\}}\right),$$

and

$$f_{n,x,a}^{(2)}(y) = \sum_{i=i_n+1}^{k_n-p} \mathbb{E}\left(g(\check{B}_n(i) - a) \mathbf{1}_{\{\check{B}_n(i) = \sqrt{b_n}\hat{B}_n(i) \leq a\}}\right).$$

Set $\phi(x) := xe^{\frac{-x^2}{x}} \mathbf{1}_{\{x \geq 0\}}$. By Theorem 1 [10] of Caravenna for $n \rightarrow \infty$,

$$\begin{aligned} & \mathbb{Q}\left(-x + y - z \leq \sqrt{b_n}\hat{B}_n(k_n - p - i) \leq -(x + y) \mid \sqrt{b_n}\hat{B}_n(j) \geq 0, j \leq k_n - p - i\right) \\ &= \frac{-z}{\sqrt{(k_n - p)b_n}} \phi\left(\frac{-y}{\sqrt{(k_n - p)b_n}}\right) + o\left(\frac{1}{\sqrt{(k_n - p)b_n}}\right), \end{aligned}$$

uniformly in $y \leq 0$, $x \in [-r_n, 0]$ and z in any compact set of \mathbb{R}_- . As a consequence by (18) we get

$$g_{x,n,y}(z) = \frac{-z}{(k_n - p)\sqrt{b_n}\pi} \phi\left(\frac{-y}{\sqrt{(k_n - p)b_n}}\right) + o\left(\frac{1}{(k_n - p)\sqrt{b_n}}\right),$$

uniformly in $y \leq 0$, $x \in [-r_n, 0]$ and $z \in [-a, 0]$. For n large enough we get

$$\begin{aligned} f_{n,x,a}^{(1)}(y) &= \frac{1}{(k_n - p)\sqrt{b_n}\pi} \phi\left(\frac{-y}{\sqrt{(k_n - p)b_n}}\right) \sum_{i=0}^{i_n} \mathbb{E}\left(-(\hat{B}_n(i) - \frac{a}{\sqrt{b_n}}) \mathbf{1}_{\{\hat{B}_n(k) \leq \frac{a}{\sqrt{b_n}}, k \leq i\}}\right) \\ &+ o\left(\frac{1}{k_n\sqrt{b_n}}\right) \sum_{i=0}^{i_n} \mathbb{Q}\left(\hat{B}_n(k) \leq \frac{a}{\sqrt{b_n}}, k \leq i\right). \end{aligned} \quad (23)$$

We now treat the quantity

$$\mathbb{E}\left(f_{x,n,a}^{(2)}(\sqrt{b_n}\hat{B}_n(k_n - p)) \mathbf{1}_{\{\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n - p\}}\right).$$

Since ϕ is bounded, there exists a constant $C > 0$ such that for all $x \in [-r_n, 0]$, $z \in \mathbb{R}$ and $0 \leq i \leq p$

$$g_{x,n,y}(z) \leq \frac{C}{\sqrt{b_n}(k_n - p - i + 1)} \mathbf{1}_{\{-a \leq z \leq 0\}},$$

as a consequence, for all $y \leq 0$ we have

$$f_{x,n,a}^{(2)}(y) \leq \frac{C}{\sqrt{b_n}} \sum_{i=i_n+1}^{k_n-p} \frac{1}{k_n - p - i + 1} \mathbb{P}\left(\check{B}_n(i) \leq \frac{a}{\sqrt{b_n}}, \hat{B}_n(i) \geq 0\right)$$

which is bounded using Lemma 7 by

$$f_{x,n,a}^{(2)}(y) \leq \frac{C}{\sqrt{b_n}} \sum_{i=i_n+1}^{k_n-p} \frac{1}{(k_n - p - i + 1)i^{\frac{3}{2}}} = o\left(\frac{1}{k_n\sqrt{b_n}}\right).$$

On the other hand we have, $\mathbb{Q}(\hat{B}_n(j) \leq \frac{-x}{\sqrt{b_n}}, j \leq k_n - p) \sim_{n \rightarrow \infty} \frac{\sqrt{2}}{\sqrt{\pi}} \frac{L(\frac{-x}{\sqrt{b_n}})}{\sqrt{k_n}}$ by (18), which mean that

$$\left(L\left(\frac{-x}{\sqrt{b_n}}\right)\right)^{-1} \mathbb{E}\left(f_{x,n,a}^{(2)}(\hat{B}_n(k_n - p)) \mathbf{1}_{\{\sqrt{b_n}\hat{B}_n(k) \leq -x, k \leq k_n - p\}}\right) = o\left(\frac{1}{k_n^{\frac{3}{2}}\sqrt{b_n}}\right). \quad (24)$$

We now return to equation (23). Letting $n \rightarrow \infty$, we have $\sum_{i=0}^{\infty} \mathbb{P}\left(\frac{\hat{T}_i}{\sqrt{b_n}} \geq \frac{-a}{\sqrt{b_n}}\right) = R\left(\frac{a}{\sqrt{b_n}}\right)$ for $a > 0$ and by Fubini's theorem we have

$$\sum_{i=0}^{\infty} \mathbb{E}\left(\left(\hat{B}_n(i) + \frac{a}{\sqrt{b_n}}\right) \mathbf{1}_{\left\{\hat{B}_n(k) \geq \frac{-a}{\sqrt{b_n}}, k \leq i\right\}}\right) = \int_0^{\frac{a}{\sqrt{b_n}}} L(t) dt = \frac{1}{\sqrt{b_n}} \int_{-a}^0 L\left(\frac{-t}{\sqrt{b_n}}\right) dt.$$

By dominated convergence theorem we have

$$\frac{1}{\sqrt{b_n}} \int_{-a}^0 L\left(\frac{-t}{\sqrt{b_n}}\right) dt = \frac{a}{\sqrt{b_n}} (1 + o(1)) = \frac{1}{\sqrt{b_n}} \int h(t) dt (1 + o(1)).$$

since $h(t) = \mathbf{1}_{[-a, 0]}(t)$. This yields, for all $y \leq 0$

$$f_{n,x,a}^{(1)}(y) = \frac{1}{k_n \sqrt{b_n} \pi} \phi\left(\frac{-y}{\sqrt{(k_n - p)b_n}}\right) \int h(t) dt + o\left(\frac{1}{k_n \sqrt{b_n}}\right).$$

as a consequence, by (18)

$$\begin{aligned} & \mathbb{E}\left(f_{x,n,a}^{(1)}(\sqrt{b_n} \hat{B}_n(k_n - p)) \mathbf{1}_{\left\{\hat{B}_n(k) \leq \frac{-x}{\sqrt{b_n}}, k \leq k_n - p\right\}}\right) = \frac{2\sqrt{2}}{\pi k_n^{\frac{3}{2}} \sqrt{b_n}} \int h(t) dt \\ & \times \mathbb{E}_{\frac{-x}{\sqrt{b_n}}}\left(\phi\left(\frac{-\hat{B}_n(k_n - p) + \frac{x}{\sqrt{b_n}}}{\sqrt{k_n - p}}\right) \mid -\hat{B}_n(k) \geq 0, k \leq k_n - p\right) \left(L\left(\frac{-x}{\sqrt{b_n}}\right) + o(1)\right). \end{aligned}$$

On the other hand, it's known (see Lemma 2.2 in [3]) that under $\mathbb{P}_y(\cdot \mid \frac{\hat{T}_k}{\sqrt{b_n}} \geq 0, k \leq k_n - p)$, $\frac{\hat{B}_n(k_n)}{\sqrt{k_n}}$ converges weakly (as $n \rightarrow \infty$) to the Rayleigh distribution with density ϕ . Hence

$$\lim_{n \rightarrow \infty} \mathbb{E}_{\frac{-x}{\sqrt{b_n}}}\left(\phi\left(\frac{\hat{B}_n(k_n - p) + \frac{x}{\sqrt{b_n}}}{\sqrt{k_n - p}}\right) \mid \hat{B}_n(k) \geq 0, k \leq k_n - p\right) = \int_0^{\infty} \phi(t)^2 dt = \frac{\sqrt{\pi}}{4}$$

uniformly in $x \in [-r_n, 0]$. Combining this with (24) we conclude that (19) holds, which allows us to complete the proof by successive approximations. \square

3.2.2 KMT coupling for random walk

We now introduce the well-known KMT Theorem [19] which is an approximation method of a random walk satisfying (H2) by a Gaussian random walk. It allows us to link estimates on random walks satisfying (H2) to the ones previously proved under assumption (H1).

Theorem 9 (Komlos-Major-Tusnàdy). *Let $(X_i)_{1 \leq i \leq n}$ be a sequence of i.i.d random variables such that $\mathbb{E}(X_i) = 0$, $0 < \mathbb{E}(X_i^2) = \sigma^2 < \infty$ and $\mathbb{E}(\exp(\theta|X_i|)) < \infty$ for some $\theta > 0$. Then there exists a sequence of i.i.d standard normal variables $(Z_i)_{1 \leq i \leq n}$ such that for all $y \geq 0$*

$$\mathbb{P}\left(\max_{1 \leq k \leq n} \left|\sigma^{-1} \sum_{i=1}^k X_i - \sum_{i=1}^k Z_i\right| > C \log(n) + y\right) \leq K \exp(-\lambda y)$$

where C, K and λ are universal positive constants.

Observe that we can construct an increasing sequence of integers d_n such that $d_n / \log n \rightarrow \infty$ and $d_n^2 / b_n \rightarrow 0$. Using assumption (6), we can apply Theorem 9, to note that with high probability, the random walk $(\bar{T}_k)_{k \leq k_n}$ stays within distance d_n from a random walk $(\hat{S}_k)_{0 \leq k \leq k_n}$ with standard Gaussian steps. More precisely, setting the event $\mathcal{W}_n = \left\{|\bar{T}_k - \sqrt{b_n} \hat{S}_k| \leq d_n, k \leq k_n\right\}$ we have

$$\mathbb{P}(\mathcal{W}_n^c) \leq e^{-\lambda(d_n - C \log n)} = o(n^{-\gamma})$$

for all $\gamma > 0$, as $d_n \gg \log n$. We start by proving a version of Lemma 4 for the random walk satisfying (H2).

Lemma 10. Fix $\epsilon > 0$, there exists $C > 0$ such that

$$\mathbb{P}(\bar{T}_k \geq -(k^{1/2-\epsilon} - y), k \leq k_n) \leq C \frac{1 + \frac{y}{\sqrt{b_n}}}{\sqrt{k_n}}$$

for any $y > 0$.

Proof. Let $\epsilon > 0$. The proof is an application Theorem 9. We have

$$\begin{aligned} & \mathbb{P}\left(\bar{T}_k \geq -(k^{1/2-\epsilon} - y), k \leq k_n\right) \\ & \leq \mathbb{P}\left(\bar{T}_k \geq -(k^{1/2-\epsilon} - y), k \leq k_n, \mathcal{W}^c, k \leq k_n\right) + \mathbb{P}\left(\sqrt{b_n}\hat{S}_k \geq -(k^{1/2-\epsilon} + d_n - y), j \leq k\right). \end{aligned}$$

Applying Theorem 9, there exists a constant $C > 0$ such that

$$\begin{aligned} & \mathbb{P}\left(\bar{T}_k \geq -(k^{1/2-\epsilon} - y), k \leq k_n\right) \\ & \leq \frac{C}{n^\gamma} + \mathbb{P}\left(\sqrt{b_n}\hat{S}_k \geq -(k^{1/2-\epsilon} + d_n - y)j \leq k\right), \end{aligned}$$

for all $\gamma > 0$. Using the fact that $\frac{d_n}{\sqrt{b_n}} \rightarrow 0$ as $n \rightarrow \infty$, we have for n large enough

$$\mathbb{P}\left(\bar{T}_k \geq -(k^{1/2-\epsilon} - y), k \leq k_n\right) \leq \mathbb{P}\left(\hat{S}_k \geq -\frac{1}{\sqrt{b_n}}(k^{1/2-\epsilon} + 1 - y)j \leq k\right)$$

and by Lemma 4 we conclude the proof. \square

In the same way we prove a similar result to Lemma 7.

Lemma 11. Let $\alpha > 0$, and for $0 \leq k \leq k_n$, write $f_n(k) = \alpha \log\left(\frac{(k_n - k)b_n + 1}{k_n b_n}\right)$. There exists $C > 0$ such that for all $x \geq 0$, $a \leq b \in \mathbb{R}$, we have for n large enough

$$\mathbb{P}\left(\bar{T}_k - f_n(k) \in [a, b], \bar{T}_j \leq f_n(j) + x, j \leq k\right) \leq C(b - a + 2d_n) \frac{(1 + \frac{x}{\sqrt{b_n}})^2}{\sqrt{b_n} k^{\frac{3}{2}}}.$$

Proof. Applying again Theorem 9, there exists a constant $C > 0$ such that

$$\begin{aligned} & \mathbb{P}\left(\bar{T}_k - f_n(k) \in [a, b], \bar{T}_j \leq f_n(j) + x, j \leq k\right) \\ & \leq \frac{C}{n^\gamma} + \mathbb{P}\left(\sqrt{b_n}\hat{S}_k - f_n(k) \in [a - d_n, b + d_n], \hat{S}_j \leq \frac{1}{\sqrt{b_n}}(f_n(j) + x + d_n), j \leq k\right). \end{aligned}$$

for all $\gamma > 0$. For n large enough we obtain

$$\begin{aligned} & \mathbb{P}\left(\bar{T}_k - f_n(k) \in [a, b], \bar{T}_j \leq f_n(j) + x, j \leq k\right) \\ & \leq \mathbb{P}\left(\sqrt{b_n}\hat{S}_k - f_n(k) \in [a - d_n, b + d_n], \hat{S}_j \leq \frac{1}{\sqrt{b_n}}(f_n(j) + x) + 2, j \leq k\right), \end{aligned}$$

and we use Lemma 7 to complete the proof. \square

We now prove a version of Lemma 8 for a random walk satisfying (\mathbf{H}_2) . Set

$$\bar{F}_n(k) = \frac{k}{k_n} a_n - c_n \mathbf{1}_{k \neq 0, k_n}$$

$k = 0, \dots, k_n$ where $(c_n)_{n \in \mathbb{N}}$ is a sequence of integers satisfying $\lim_{n \rightarrow \infty} c_n = \infty$ and

$$\lim_{n \rightarrow \infty} \frac{c_n}{\sqrt{b_n}} = 0. \quad (25)$$

Before moving to the proof we need the following Lemma.

Lemma 12. Uniformly in $x \in [c_n, \infty]$, we have

$$\lim_{n \rightarrow \infty} \left| \frac{L\left(\frac{x - c_n}{\sqrt{b_n}}\right)}{L\left(\frac{x}{\sqrt{b_n}}\right)} - 1 \right| = 0.$$

Proof. The proof follows from the properties of the renewal function introduced in (14). We first consider the ratio $\frac{L(y - \frac{c_n}{\sqrt{b_n}})}{L(y)}$ for large value of y . Let $\epsilon > 0$, by (15), there exists a constant $A = A(\epsilon) > 0$ sufficiently large such that for all $y \geq A$ we have

$$c_0(1 - \epsilon)y \leq L(y) \leq c_0(1 + \epsilon)y. \quad (26)$$

Using (26) we have

$$\sup_{\frac{x}{\sqrt{b_n}} \geq A} \left| \frac{L(\frac{x - c_n}{\sqrt{b_n}})}{L(\frac{x}{\sqrt{b_n}})} - 1 \right| \leq \left| \frac{1 + \epsilon}{1 - \epsilon} - 1 \right| + \frac{c_0(1 + \epsilon)c_n}{A\sqrt{b_n}}. \quad (27)$$

On the other hand, recall that $y \mapsto L(y)$ is continuous. Hence it is uniformly continuous on $[0, A]$, and

$$\begin{aligned} & \sup_{x \in [c_n, A\sqrt{b_n}]} \left| L\left(\frac{x}{\sqrt{b_n}}\right) - L\left(\frac{x - c_n}{\sqrt{b_n}}\right) \right| \\ & \leq \sup_{y \in [\frac{c_n}{\sqrt{b_n}}, A]} \left| L(y) - L\left(y - \frac{c_n}{\sqrt{b_n}}\right) \right| \leq w_L\left(\frac{c_n}{\sqrt{b_n}}\right), \end{aligned}$$

where $w_L(\delta) = \sup_{\substack{s, t \\ |t - s| \leq \delta}} |L(t) - L(s)|$. Since the renewal function is increasing, we have $L(y) \geq 1$, for all $y \geq 0$, which implies that

$$\sup_{y \in [\frac{c_n}{\sqrt{b_n}}, A]} \left| \frac{L(y - \frac{c_n}{\sqrt{b_n}})}{L(y)} - 1 \right| \leq w_L\left(\frac{c_n}{\sqrt{b_n}}\right) \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (28)$$

By (27) and (28) we have

$$\begin{aligned} \sup_{x \in [c_n, \infty)} \left| \frac{L(\frac{x - c_n}{\sqrt{b_n}})}{L(\frac{x}{\sqrt{b_n}})} - 1 \right| & \leq \sup_{x \in [c_n, A\sqrt{b_n}]} \left| \frac{L(\frac{x - c_n}{\sqrt{b_n}})}{L(\frac{x}{\sqrt{b_n}})} - 1 \right| + \sup_{\frac{x}{\sqrt{b_n}} \geq A} \left| \frac{L(\frac{x - c_n}{\sqrt{b_n}})}{L(\frac{x}{\sqrt{b_n}})} - 1 \right| \\ & \leq w_L\left(\frac{c_n}{\sqrt{b_n}}\right) + \left| \frac{1 + \epsilon}{1 - \epsilon} - 1 \right| + \frac{c_0(1 + \epsilon)c_n}{A\sqrt{b_n}}. \end{aligned}$$

Letting $n \rightarrow \infty$ then $\epsilon \rightarrow 0$ we conclude the proof. \square

Corollary 13. *Setting*

$$a_n = \frac{-3}{2\theta^*} \ln(n) + \frac{\log(b_n)}{\theta^*},$$

then there exists $C > 0$ such that With the same notation as Lemma 8 we have

$$\mathbb{E}(f(\bar{T}_{k_n} - a_n + x)e^{-\theta^* \bar{T}_{k_n}} \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}}) = \frac{e^{\theta^* x}}{\sqrt{2\pi\sigma^2 b_n}} \int_{-\infty}^0 f(y)e^{-\theta^* y} dy \left(L\left(\frac{-x}{\sqrt{b_n}}\right) + o(1) \right), \quad (29)$$

uniformly in $x \in [-r_n, -(c_n + d_n)]$.

Proof. We will divide the proof of this corollary into two parts, proving separately an upper and a lower bound for (29). We set $h(z) = e^{-\theta^* z} f(z)$. As a consequence, it is enough to prove that

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}}) = \frac{1}{k_n^{3/2} \sqrt{2\pi\sigma^2 b_n}} \int_{-\infty}^0 h(y) dy \left(L\left(\frac{-x}{\sqrt{b_n}}\right) + o(1) \right)$$

uniformly in $x \in [-r_n, -(c_n + d_n)]$. Using the same arguments in Lemma 8, it is enough to prove this Corollary with the function $h(z) = \mathbf{1}_{[-a, 0]}(z)$ for some $a > 0$. Then we write

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}}) = \mathbb{P}(\mathcal{A}_n^{(k)}(x)) \quad (30)$$

where

$$\mathcal{A}_n^{(k)}(x) = \{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n, \bar{T}_{k_n} - a_n + x \geq -a\}$$

the event that the random walk $(\bar{T}_k)_{k \leq k_n}$ stays below the barrier $k \mapsto \bar{F}_n(k)$ for all $k \leq k_n$ and end up in a finite interval.

• **Upper bound**

We have

$$\mathbb{P}(\mathcal{A}_n^{(k)}(x)) = \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), |\bar{T}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n}\right) + \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), |\bar{T}_{k_n-1} - a_n + x| > h_n \sqrt{b_n}\right) \quad (31)$$

where $(h_n)_{n \in \mathbb{N}}$ is a sequence growing to ∞ , that we will fix later on. We bound these two quantities separately choosing h_n such that

$$\limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} k_n^{3/2} \sqrt{b_n} \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), |\bar{T}_{k_n-1} - a_n + x| > h_n \sqrt{b_n}\right) = 0.$$

We first observe that

$$\mathbb{P}\left(\mathcal{A}_n^{(k)}(x), |\bar{T}_{k_n-1} - a_n + x| > h_n \sqrt{b_n}\right) = \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), \bar{T}_{k_n-1} - a_n + x < -h_n \sqrt{b_n}\right)$$

since the event $\{\bar{T}_{k_n-1} \leq \bar{F}_n(k_n-1) - x, \bar{T}_{k_n-1} - a_n + x > h_n \sqrt{b_n}\}$ is impossible for n large enough. Then we have,

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} k_n^{3/2} \sqrt{b_n} \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), \bar{T}_{k_n-1} - a_n + x < -h_n \sqrt{b_n}\right) \\ & \leq \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} k_n^{3/2} \sqrt{b_n} \mathbb{P}\left(\bar{T}_{k_n} - a_n + x \in [-a, 0], \bar{T}_{k_n-1} - a_n + x < -h_n \sqrt{b_n}\right) \\ & \leq \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} k_n^{3/2} \sqrt{b_n} \mathbb{P}(|\bar{T}_1| \geq \sqrt{b_n} h_n). \end{aligned}$$

On the other hand we write

$$\mathbb{P}(|\bar{T}_1| \geq \sqrt{b_n} h_n) = \mathbb{P}(\bar{T}_1 \geq \sqrt{b_n} h_n) + \mathbb{P}(-\bar{T}_1 \geq \sqrt{b_n} h_n),$$

and additionally we have

$$\mathbb{P}\left(\bar{T}_1 \geq \sqrt{b_n} h_n\right) = \mathbb{P}\left(e^{\theta \bar{T}_1} \geq e^{\theta h_n \sqrt{b_n}}\right) \leq e^{b_n(\Lambda(\theta) - \theta h_n \sqrt{b_n})},$$

for all $\theta > 0$. As Λ is \mathcal{C}^2 on a neighbourhood of 0, we have

$$\Lambda(\theta) = \Lambda(0) + \Lambda'(0)\theta + \frac{\Lambda''(0)\theta^2}{2} + o(\theta^2) = \Lambda''(0)\theta^2/2 + o(\theta^2).$$

Therefore, choosing $\theta = \frac{h_n}{\sqrt{b_n}}$, there exists $C > 0$ such that for all n large enough, we have

$$\mathbb{P}(\bar{T}_1 \geq \sqrt{b_n} h_n) \leq e^{-Ch_n^2}.$$

Similarly, we have $\mathbb{P}(\bar{T}_1 \leq -\sqrt{b_n} h_n) \leq e^{-Ch_n^2}$ for n large enough. Finally we obtain

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} k_n^{3/2} \sqrt{b_n} \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), \bar{T}_{k_n-1} - a_n + x < -h_n \sqrt{b_n}\right) \\ & \leq \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} k_n^{3/2} \sqrt{b_n} e^{-Ch_n^2} \end{aligned}$$

which goes to zero as $n \rightarrow \infty$ as long as $h_n > 2\sqrt{\frac{\log(n)}{C}}$.

We now bound the first quantity in the right hand-side of (31). Applying the Markov property at time $k_n - 1$ we get

$$\begin{aligned} & \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), |\bar{T}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n}\right) \\ & = \mathbb{E}\left(f_n(\bar{T}_j + a_n - x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, |\bar{T}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n}, k \leq k_n - 1\}}\right) \end{aligned} \quad (32)$$

where $f_n(z) = \mathbb{P}_z(\bar{T}_1 \in [-a, 0])$ for all $z \in \mathbb{R}$. We estimate the function $z \mapsto f_n(z)$, $n \in \mathbb{N}$, using the refined Stone's local limit theorem in [8, Theorem 2.1]. By (\mathbf{H}_2) there exists a constant $c > 0$ such that for all $z \in \mathbb{R}$

$$f_n(z) \leq \frac{a}{\sqrt{2\pi b_n \sigma^2}} \exp\left(\frac{-(z - a_n + x)^2}{2b_n}\right) + \frac{c}{b_n}.$$

To approximate (32) for a random walk satisfying (\mathbf{H}_2) we apply Theorem 9, there exists a constant $C > 0$ such that for all $\gamma > 0$ we have

$$\begin{aligned} & \mathbb{P}\left(\mathcal{A}_n^{(k)}(x), |\bar{T}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n}\right) \\ & \leq \frac{C}{n^\gamma} + \frac{a}{\sqrt{2\pi b_n \sigma^2}} \mathbb{E}\left(\sup_{|w| \leq d_n} \exp\left(\frac{-(\hat{S}_{k_n-1} + w - a_n + x)^2}{2b_n \sigma^2}\right) \mathbf{1}_{\mathcal{B}_n^{(k)}(x)}\right) + \frac{c}{b_n} \mathbb{P}\left(\mathcal{B}_n^{(k)}(x)\right). \end{aligned} \quad (33)$$

where $\mathcal{B}_n^{(k)}(x) = \left\{ \hat{S}_k \leq \tilde{F}_n(k) + d_n - x, k \leq k_n - 1, |\hat{S}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n} + d_n \right\}$. We write then

$$\begin{aligned} & \mathbb{E}\left(\sup_{|w| \leq d_n} \exp\left(\frac{-(\hat{S}_{k_n-1} + x - a_n + w)^2}{2b_n}\right) \mathbf{1}_{\mathcal{B}_n^{(k)}(x)}\right) \\ & = \mathbb{E}_{\mathbb{Q}}\left(\sup_{|w| \leq d_n} e^{-\frac{a_n}{n} \sqrt{b_n} \tilde{S}_{k_n-1} + n\Lambda(a_n/n)} \exp\left(\frac{-(\sqrt{b_n} \tilde{S}_{k_n-1} + x - a_n + w)^2}{2b_n}\right) \mathbf{1}_{\mathcal{C}_n^{(k)}(x)}\right) \end{aligned} \quad (34)$$

where

$$\mathcal{C}_n^{(k)}(x) = \left\{ \tilde{S}_k \leq \frac{-x + d_n - c_n}{\sqrt{b_n}}, k \leq k_n - 1, |\sqrt{b_n} \tilde{S}_{k_n-1} + \frac{a_n}{k_n} - x| \leq h_n \sqrt{b_n} + d_n \right\}$$

and $\sqrt{b_n} \tilde{S}_k = \sqrt{b_n} \hat{S}_k - \frac{k}{k_n} a_n$ is a centred Gaussian random walk under the measure \mathbb{Q} defined in the proof of Lemma 8. Then using the dominated convergence theorem and the fact that $\frac{d_n^2}{b_n} \rightarrow_{n \rightarrow \infty} 0$, and rescaling by $\frac{1}{\sqrt{b_n}}$ we obtain

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \mathbb{E}\left(\exp\left(\sup_{|w| \leq d_n} \frac{-(\hat{S}_{k_n-1} + w - a_n + x)^2}{2b_n \sigma^2}\right) \mathbf{1}_{\mathcal{C}_n^{(k)}(x)}\right) \\ & \leq \limsup_{n \rightarrow \infty} \mathbb{E}_{\mathbb{Q}}\left(\exp\left(\frac{-(\sqrt{b_n} \tilde{S}_{k_n-1} + x)^2}{2b_n \sigma^2}\right) \mathbf{1}_{\left\{\tilde{S}_k \leq \frac{1}{\sqrt{b_n}}(x + d_n - c_n), k \leq k_n - 1\right\}}\right) \\ & \leq \limsup_{n \rightarrow \infty} \mathbb{E}_{\mathbb{Q}}\left(\exp\left(\frac{-(\tilde{S}_{k_n-1} + \frac{-x}{\sqrt{b_n}})^2}{2\sigma^2}\right) \mathbf{1}_{\left\{\tilde{S}_k \leq \frac{-x + d_n - c_n}{\sqrt{b_n}}, k \leq k_n - 1\right\}}\right). \end{aligned}$$

Applying Lemma 2.3 in [2] for the Gaussian random walk $(\tilde{S}_k)_{k \leq k_n}$, and using Lemma 12 we obtain for all $\frac{x}{\sqrt{b_n}} \in \left[-\frac{r_n}{\sqrt{b_n}}, -\frac{(c_n + d_n)}{\sqrt{b_n}}\right]$,

$$\begin{aligned} & \limsup_{n \rightarrow \infty} k_n^{\frac{3}{2}} \mathbb{E}_{\mathbb{Q}}\left(\exp\left(\frac{-(\tilde{S}_{k_n-1} + \frac{-x}{\sqrt{b_n}})^2}{2\sigma^2}\right) \mathbf{1}_{\left\{\tilde{S}_k \leq \frac{-x + d_n - c_n}{\sqrt{b_n}}, k \leq k_n - 1\right\}}\right) \\ & \leq \frac{1}{\sqrt{2\pi\sigma^2}} \int_0^\infty e^{-\frac{y^2}{2\sigma^2}} L(y) dy \times L\left(\frac{-x + d_n - c_n}{\sqrt{b_n}}\right). \end{aligned}$$

We now observe that that $\frac{1}{\sqrt{2\pi\sigma^2}} \int_0^\infty L(y) e^{-\frac{y^2}{2\sigma^2}} dy = \mathbb{E}\left(L(\hat{S}_1) \mathbf{1}_{\{\hat{S}_1 \geq 0\}}\right) = L(0) = 1$, which implies

$$\begin{aligned} & \limsup_{n \rightarrow \infty} k_n^{\frac{3}{2}} \mathbb{E}_{\mathbb{Q}}\left(\exp\left(\frac{-(\sqrt{b_n} \tilde{S}_{k_n-1} + x)^2}{2b_n \sigma^2}\right) \mathbf{1}_{\left\{\tilde{S}_k \leq \frac{-x + d_n - c_n}{\sqrt{b_n}}, k \leq k_n - 1\right\}}\right) \\ & \leq L\left(\frac{-x + d_n - c_n}{\sqrt{b_n}}\right). \end{aligned}$$

To complete the proof of the upper bound it remains to show that

$$\limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} k_n^{3/2} b_n^{-\frac{1}{2}} \mathbb{E} \left(\mathbf{1}_{\mathcal{B}_n^k(x)} \right) = 0. \quad (35)$$

Using similar computations we have

$$\begin{aligned} & \sup_{x \in [-r_n, -c_n]} k_n^{\frac{3}{2}} \mathbb{P} \left(\hat{S}_k \leq \tilde{F}_n(k) + d_n - x, k \leq k_n - 1, |\hat{S}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n} + d_n \right) \\ & \leq \sup_{x \in [-r_n, -c_n]} k_n^{\frac{3}{2}} \mathbb{P} \left(\tilde{S}_k \leq \frac{-x + d_n - c_n}{\sqrt{b_n}}, \tilde{S}_{k_n-1} - \frac{a_n}{k_n \sqrt{b_n}} + \frac{x}{\sqrt{b_n}} \geq -h_n - \frac{d_n}{\sqrt{b_n}}, k \leq k_n - 1 \right) \\ & \leq CL \left(\frac{-x + d_n - c_n}{\sqrt{b_n}} \right) \int_0^{h_n + \frac{2d_n}{\sqrt{b_n}}} L(y) dy. \end{aligned}$$

Since the renewal function $x \mapsto R(x)$ is increasing we have by (17)

$$\begin{aligned} & \sup_{x \in [-r_n, -c_n]} k_n^{\frac{3}{2}} \mathbb{P} \left(\hat{S}_k \leq \tilde{F}_n(k) + d_n - x, k \leq k_n - 1, |\hat{S}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n} + d_n \right) \\ & \leq C \left(h_n + 2 \frac{d_n}{\sqrt{b_n}} \right) L \left(\frac{-x + d_n - c_n}{\sqrt{b_n}} \right) L \left(h_n + \frac{d_n}{\sqrt{b_n}} \right) \leq C \left(h_n + \frac{2d_n}{\sqrt{b_n}} \right)^2 L \left(\frac{-x + d_n - c_n}{\sqrt{b_n}} \right) \end{aligned}$$

where we used in the last inequality that $L(t_n) \leq ct_n$ for some constant $c > 0$ when $t_n \rightarrow_{n \rightarrow \infty} \infty$. Thanks to Lemma 12 we obtain

$$\sup_{x \in [-r_n, -c_n]} L \left(\frac{-x}{\sqrt{b_n}} \right)^{-1} k_n^{3/2} b_n^{-\frac{1}{2}} \mathbb{P} \left(\mathcal{B}_n^{(k)}(x) \right) \leq C \frac{\left(h_n + \frac{2d_n}{\sqrt{b_n}} \right)^2}{\sqrt{b_n}}$$

which goes to zero since $\frac{\log(n) + d_n}{\sqrt{b_n}} \rightarrow 0$ as $n \rightarrow \infty$. Finally by (33), (35) and Lemma 12 we deduce that

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L \left(\frac{-x}{\sqrt{b_n}} \right)^{-1} \sqrt{b_n} k_n^{\frac{3}{2}} \mathbb{E} \left(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}} \right) \\ & \leq \limsup_{n \rightarrow \infty} \sup_{x \in [-r_n, -c_n]} L \left(\frac{-x}{\sqrt{b_n}} \right)^{-1} \sqrt{b_n} k_n^{\frac{3}{2}} \mathbb{P} \left(\mathcal{A}_n^{(k)}(x), |\bar{T}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n} \right) \\ & \leq \frac{1}{\sqrt{2\sigma^2\pi}} \int_{-\infty}^0 f(y) e^{-\theta^* y} dy. \end{aligned} \quad (36)$$

We now treat the lower bound.

- **Lower bound**

We now compute a lower bound for $\mathbb{P} \left(\mathcal{A}_n^{(k)}(x) \right)$. Using similar arguments to these used in (32) by ignoring in this case the event $\{|\bar{T}_{k_n-1} - a_n + x| \leq h_n \sqrt{b_n}\}$, which does not play any role in the proof of the lower bound we have by Theorem 9

$$\begin{aligned} & \mathbb{P} \left(\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n, \bar{T}_{k_n} - a_n + x \in [-a, 0] \right) \\ & \geq \frac{a}{\sqrt{2\pi b_n \sigma^2}} \mathbb{E} \left(\inf_{|w| \leq d_n} \exp \left(\frac{-(\sqrt{b_n} \hat{S}_{k_n-1} + w - a_n + x)^2}{2b_n \sigma^2} \right) \mathbf{1}_{\{\sqrt{b_n} \hat{S}_k \leq \bar{F}_n(k) - d_n - x, k \leq k_n - 1\}} \right). \end{aligned}$$

Similar computations to these used in the upper bound lead to

$$\begin{aligned} & \liminf_{n \rightarrow \infty} \mathbb{E} \left(\inf_{|w| \leq d_n} \exp \left(\frac{-(\sqrt{b_n} \hat{S}_{k_n-1} + x - \frac{a_n}{n} + w)^2}{2b_n} \right) \mathbf{1}_{\{\hat{S}_k \leq \bar{F}_n(k) - x - d_n, k \leq k_n - 1\}} \right) \\ & \geq \liminf_{n \rightarrow \infty} \mathbb{E}_{\mathbb{Q}} \left(\exp \left(\frac{-(\tilde{S}_{k_n-1} + \frac{x}{\sqrt{b_n}})^2}{2\sigma^2} \right) \mathbf{1}_{\{\tilde{S}_k \leq \frac{-x - d_n - c_n}{\sqrt{b_n}}, k \leq k_n - 1\}} \right). \end{aligned}$$

Applying again Lemma 2.3 in [2] for the Gaussian random walk $(\tilde{S}_k)_{k \leq k_n}$, we obtain for all $x \in [-r_n, -(c_n + d_n)]$,

$$\begin{aligned} & \liminf_{n \rightarrow \infty} k_n^{\frac{3}{2}} \mathbb{E}_{\mathbb{Q}} \left(\exp \left(\frac{-(\tilde{S}_{k_n-1} + \frac{x}{\sqrt{b_n}})^2}{2\sigma^2} \right) \mathbf{1}_{\left\{ \tilde{S}_k \leq \frac{-x - d_n - c_n}{\sqrt{b_n}}, k \leq k_n - 1 \right\}} \right) \\ & \geq \frac{1}{\sqrt{2\pi\sigma^2}} \int_0^\infty e^{-\frac{y^2}{2\sigma^2}} R(y) dy \times L\left(\frac{-x - (c_n + d_n)}{\sqrt{b_n}}\right). \end{aligned}$$

Finally, by Lemma 12 we get

$$\begin{aligned} & \liminf_{n \rightarrow \infty} \inf_{x \in [-r_n, -(c_n + d_n)]} \left(L\left(\frac{-x}{\sqrt{b_n}}\right)^{-1} \sqrt{b_n} k_n^{\frac{3}{2}} \mathbb{E}(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}}) \right) \\ & \geq \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^0 f(y) e^{-\theta^* y} dy. \end{aligned} \quad (37)$$

Combining equations (36) and (37) we deduce that

$$\mathbb{E}(h(\bar{T}_{k_n} - a_n + x) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k) - x, k \leq k_n\}}) = \frac{e^{\theta^* x}}{\sqrt{2\pi\sigma^2} \sqrt{b_n} k_n^{\frac{3}{2}}} \int_{-\infty}^0 f(y) e^{-\theta^* y} dy \left(L\left(\frac{-x}{\sqrt{b_n}}\right) + o(1) \right)$$

uniformly in $x \in [-r_n, -(c_n + d_n)]$. \square

4 The modified extremal process

Recall that $(\bar{T}_k)_{1 \leq k \leq k_n}$ is a sequence of centred random walk with $\text{Var}(\bar{T}_k) = kb_n\sigma^2$ and $a_n = m_n - k_n b_n v$. The goal of the next section is to introduce a modified extremal process and to prove that it has the same weak limit as the original extremal process \mathcal{E}_n .

Start by setting

$$\mathcal{E}_{n, R_n} = \sum_{u \in \mathcal{H}_{k_n}} \delta_{S_u - m_n} \mathbf{1}_{\{S_{u_k} \leq R_n(k), \forall k \leq k_n\}},$$

where refer to the function $R_n : \{0, \dots, k_n\} \mapsto \mathbb{R}$ as a barrier. More precisely our objective is to prove that the weak limit of the modified extremal process \mathcal{E}_{n, R_n} and the original extremal process \mathcal{E}_n coincide for a well-chosen function R_n .

In a first step, we start by proving that there exists a barrier R_n such that, with high probability, all individuals stay below it all most of time. The existence of this barrier gives us further information for the localization of the paths of the extremal particles.

Lemma 14. *Consider the barrier*

$$R_n(k) = kb_n v - \frac{3}{2\theta^*} \log\left(\frac{k_n b_n + 1}{(k_n - k)b_n + 1}\right) + c_n, k = 0, \dots, k_n$$

where $(c_n)_{n \in \mathbb{N}}$ is the sequence of integers defined in (25). It then holds:

$$\mathbb{P}(\exists u \in \mathcal{H}_{k_n}, S_{u_k} > R_n(k), \text{ for some } k \leq k_n) = o(1) \text{ when } n \uparrow \infty.$$

Proof. Using Markov inequality we get

$$\mathbb{P}(\exists |u| = k_n, S_{u_k} > R_n(k), k \leq k_n) \leq \sum_{k \leq k_n} \mathbb{E} \left(\sum_{|u|=k} \mathbf{1}_{\{S_{u_k} > R_n(k), S_{u_j} \leq R_n(j), j < k\}} \right).$$

By Proposition 3 we have,

$$\begin{aligned} & \sum_{k \leq k_n} \mathbb{E} \left(\sum_{|u|=k} \mathbf{1}_{\{S_{u_k} > R_n(k), S_{u_j} \leq R_n(j), j < k\}} \right) \\ & \leq \sum_{k \leq k_n} \mathbb{E} \left(\exp(-\theta^* \bar{T}_k) \mathbf{1}_{\{\bar{T}_k > R_n(k) - kb_n v, \bar{T}_j \leq R_n(j) - j b_n v, j < k\}} \right) \\ & \leq e^{-\theta^* c_n} \sum_{k \leq k_n} \frac{(k_n b_n + 1)^{\frac{3}{2}}}{(k_n - k)b_n + 1)^{\frac{3}{2}}} \mathbb{P}(\bar{T}_k > \bar{R}_n(k), \bar{T}_j \leq \bar{R}_n(j), j < k), \end{aligned} \quad (38)$$

where $\bar{R}_n(j) = R_n(j) - j b_n v$, for all $1 \leq j \leq k_n$. We compute this probability by conditioning with respect to the last step $\bar{T}_k - \bar{T}_{k-1}$ to get

$$\mathbb{P}(\bar{T}_k > \bar{R}_n(k), \bar{T}_j \leq \bar{R}_n(j), j \leq k) = \mathbb{E}(f_{k-1}(\bar{T}_k - \bar{T}_{k-1}))$$

where, $\forall y \in \mathbb{R}$

$$f_{k-1}(y) = \mathbb{P}\left(\bar{R}_n(k) - y \leq \bar{T}_{k-1} \leq \bar{R}_n(k), B_n(j) \leq \frac{1}{\sqrt{b_n}} \bar{R}_n(j), j \leq k-1\right).$$

Assume that (\mathbf{H}_1) or (\mathbf{H}_2) hold, by Lemma 7 or 11 and using the fact that $\frac{c_n}{\sqrt{b_n}} \rightarrow_{n \rightarrow \infty} 0$, we deduce that, for n large enough, there exists $C > 0$ such that, $\forall y \in \mathbb{R}$

$$f_{k-1}(y) \leq C \mathbf{1}_{\{y \geq 0\}} \frac{(1 + \frac{y}{\sqrt{b_n}})^3}{k^{3/2}}.$$

Now Plugging this in (38) we obtain

$$\begin{aligned} & \mathbb{P}(\exists |u| = k_n, S_{u_k} > R_n(k), k \leq k_n) \\ & \leq C e^{-\theta^* c_n} \sum_{k \leq k_n} \frac{(k_n b_n + 1)^{\frac{3}{2}}}{((k_n - k) b_n + 1)^{\frac{3}{2}}} \frac{1}{k^{\frac{3}{2}}} (1 + \mathbb{E}(\frac{\bar{T}_k - \bar{T}_{k-1}}{\sqrt{b_n}})_+^3) \rightarrow_{n \rightarrow \infty} 0. \end{aligned}$$

completing the proof. \square

From this lemma we deduce that the extremal process \mathcal{E}_{n, R_n} has the same weak limit of the one of \mathcal{E}_n . Now we will consider a barrier which is lower. For the choice of such barrier we refer to the work of Arguin, Bovier and Kistler [5] in the case of branching Brownian motion. Their method is based on the mechanism of entropic repulsion which is used in the statistical mechanics of membrane models, see [26].

Proposition 15. *Define the barrier*

$$F_n(k) = k b_n v + \frac{k}{k_n} a_n - c_n \mathbf{1}_{k \neq 0, k_n}, k = 0, \dots, k_n.$$

Let $A = [a, \infty)$ where $a \in \mathbb{R}$, then we have

$$\lim_{n \rightarrow \infty} \mathbb{E}(\mathcal{E}_{n, R_n}(A) - \mathcal{E}_{n, F_n}(A)) = 0.$$

Proof. Let the following subsets of \mathcal{H}_{k_n}

$$\mathbf{A}_n^{(u)} = \{u \in \mathcal{H}_{k_n} : S_{k_n} - m_n \in A, S_k \leq R_n(k), k \leq k_n, \}$$

the set of particles at generation k_n that are close to the maximum and that stay below the barrier $k \mapsto R_n(k)$ for all $k \leq k_n$. Respectively we introduce

$$\mathbf{B}_n^{(u)} = \{u \in \mathcal{H}_{k_n} : S_{k_n} - m_n \in A, S_k \leq F_n(k), k \leq k_n, \}$$

Set the integer-valued variable

$$\#(\mathbf{A}_n^{(u)} \cap (\mathbf{B}_n^{(u)})^c) = \#\{u \in \mathcal{H}_{k_n} : S_{k_n} - m_n \in A, S_k \leq R_n(k), k \leq k_n, \exists j \leq k_n, S_j > F_n(j)\}.$$

Using the fact that $\mathbf{B}_n^{(u)} \subset \mathbf{A}_n^{(u)}$ we deduce that

$$\begin{aligned} & \mathbb{E}(\mathcal{E}_{n, R_n}(A) - \mathcal{E}_{n, F_n}(A)) = \mathbb{E}(\#(\mathbf{A}_n^{(u)} \cap (\mathbf{B}_n^{(u)})^c)) \\ & = \mathbb{E}\left(\sum_{|u|=k_n} \mathbf{1}_{\{S_{k_n} - m_n \in A, S_k \leq R_n(k), k \leq k_n, \exists j \leq k_n, S_j > F_n(j)\}}\right), \end{aligned}$$

then thanks to Proposition 3 we obtain

$$\begin{aligned} & \mathbb{E}\left(\sum_{|u|=k_n} \mathbf{1}_{\{S_{k_n} - m_n \in A, S_k \leq R_n(k), k \leq k_n, \exists j \leq k_n, S_j > F_n(j)\}}\right) \\ & = \mathbb{E}\left(e^{-\theta^* \bar{T}_{k_n}} \mathbf{1}_{\{\bar{T}_{k_n} - a_n \in A, \bar{T}_k \leq \bar{R}_n(k), k \leq k_n, \exists j \leq k_n, \bar{T}_j > \bar{F}_n(j)\}}\right), \end{aligned}$$

where $a_n = m_n - k_n b_n v$ and $\bar{F}_n(j) = \frac{j}{k_n} a_n - c_n \mathbf{1}_{j \neq 0, k_n}$ for all $1 \leq j \leq k_n$. Summing with respect to the value of $T_{k_n} - a_n - a$ at time k_n , we have

$$\begin{aligned} & \mathbb{E} \left(e^{-\theta^* \bar{T}_{k_n}} \mathbf{1}_{\{\bar{T}_{k_n} - a_n \in A, \bar{T}_k \leq R_n(k) - k b_n v, k \leq k_n, \exists j \leq k_n, \bar{T}_j > \bar{F}_n(j)\}} \right) \\ & \leq \frac{C n^{\frac{3}{2}}}{b_n} \sum_{r \geq 0} e^{-\theta^* r} \sum_{0 \leq j \leq k_n} \mathbb{P}(\bar{T}_{k_n} - a_n - a \in [r, r+1], \bar{T}_k \leq \bar{R}_n(k), k \leq k_n, \bar{T}_j > \bar{F}_n(j)). \end{aligned}$$

Applying the Markov property at time j we get

$$\begin{aligned} & \mathbb{P}(\bar{T}_{k_n} - a_n - a \in [r, r+1], \bar{T}_k \leq \bar{R}_n(k), k \leq k_n, \bar{T}_j > \bar{F}_n(j)) \\ & \leq \mathbb{P}(\bar{T}_k \leq \bar{R}_n(k), k \leq j, \bar{T}_j > \bar{F}_n(j)) \times \end{aligned} \quad (39)$$

$$\sup_{x \in [F_n(j), R_n(j)]} \mathbb{P}_x(\bar{T}_{k_n-j} - a_n - a \in [r, r+1], \bar{T}_k \leq \bar{R}_n(k+j), k \leq k_n - j). \quad (40)$$

To bound the probability (39), we apply the Markov property at time $l = \lfloor \frac{j}{3} \rfloor$,

$$\begin{aligned} & \mathbb{P}(\bar{T}_j > \bar{F}_n(j), \bar{T}_k \leq \bar{R}_n(k), k \leq j) \\ & \leq \mathbb{P}(\bar{T}_k \leq \bar{R}_n(k), k \leq l) \sup_{z \leq \bar{R}_n(l)} \mathbb{P}_z(\bar{T}_{j-l} > \bar{F}_n(j), \bar{T}_k \leq \bar{R}_n(k+l), k \leq j-l). \end{aligned}$$

Set $\widehat{T}_k = \bar{T}_{j-l} - \bar{T}_{j-l-k}$, which is a random walk with the same law as \bar{T}_k . Then we obtain

$$\begin{aligned} & \mathbb{P}_z(\bar{T}_{j-l} > \bar{F}_n(j), \bar{T}_k \leq \bar{R}_n(k+l), k \leq j-l) \\ & \leq \mathbb{P}_z(\bar{F}_n(j) < \bar{T}_{j-l} \leq \bar{R}_n(j), \bar{T}_k \geq F_n(j) - R_n(j-k), k \leq j-l). \end{aligned} \quad (41)$$

We bound the probability in (41). We use a lower bound for the expression $(F_n(j) - R_n(j-k), k \leq j-l)$. Observe that the function $x \mapsto \frac{\log(x)}{x}$ is decreasing for $x \geq e$, and

$$(k_n - j)b_n + 1 + k b_n \leq 2((k_n - j)b_n + 1)k b_n,$$

then we have

$$\begin{aligned} F_n(j) - R_n(j-k) &= \frac{-3}{2\theta^*} \left(\frac{j}{k_n} \log(k_n b_n) - \log(k_n b_n) + \log((k_n - j + k)b_n + 1) \right) + \frac{\ln(b_n)}{\theta^*} - 2c_n \\ &\geq \frac{-3}{2\theta^*} (\log((j \vee e)b_n) - \log(k_n b_n) + \log((k_n - j)b_n + 1) + \log(k b_n) + \log(2)) - 2c_n \\ &\geq \frac{-3}{2\theta^*} (\log((j \vee e) \wedge ((k_n - j) + 1)) + \log(k b_n) + \log(2)) - 2c_n, \forall k, j = 1 \dots k_n. \end{aligned}$$

Applying again the Markov property at time l we get

$$\begin{aligned} & \mathbb{P}_z(\bar{F}_n(j) < \bar{T}_{j-l} \leq \bar{R}_n(j), \bar{T}_k \geq F_n(j) - R_n(j-k), k \leq j-l) \\ & \leq \mathbb{P} \left(\bar{T}_k \geq \frac{-3}{2\theta^*} (\log(k b_n) + \log(j \wedge ((k_n - j) + 1)) + \log(2)) - 2c_n, k \leq l \right) \times \\ & \mathbb{P}_x(\bar{F}_n(j) - z < \bar{T}_{j-2l} \leq \bar{R}_n(j) - z). \end{aligned}$$

Assume that **(H₁)** or **(H₂)** hold, then by Lemmas 4, 5 or 10 for n large enough we obtain

$$\begin{aligned} & \mathbb{P}(\bar{T}_k \leq \bar{R}_n(k), k \leq j, \bar{T}_j > \bar{F}_n(j)) \\ & \leq \underbrace{\mathbb{P} \left(B_n(k) \leq \frac{1}{\sqrt{b_n}} \bar{R}_n(k), k \leq l \right)}_{\leq \frac{C}{\sqrt{j}}} \times \\ & \underbrace{\mathbb{P} \left(B_n(k) \geq \frac{-3}{2\theta^*} (\log(j \wedge ((k_n - j) + 1)) + \log(k b_n) + \log(2)) - 2 \frac{c_n}{\sqrt{b_n}}, k \leq l \right)}_{\leq C \frac{1 + \log(j \wedge ((k_n - j) + 1))}{\sqrt{j}}} \times \\ & \underbrace{\mathbb{P}_x \left(\frac{1}{\sqrt{b_n}} (\bar{F}_n(j) - z) < B_n(j-2l) \leq \frac{1}{\sqrt{b_n}} (\bar{R}_n(j) - z) \right)}_{\leq C \frac{\log(j \wedge ((k_n - j) + 1)) + 1}{\sqrt{j b_n}}} \leq C \frac{(1 + (\log(j \wedge ((k_n - j) + 1)))^2}{\sqrt{b_n} j^{\frac{3}{2}}}. \end{aligned}$$

Using Lemma 7 or 11, for n large enough we have

$$\begin{aligned} & \sup_{x \in [F_n(j), R_n(j)]} \mathbb{P}_x (\bar{T}_{k_n-j} - a_n - a \in [r, r+1], \bar{T}_k \leq \bar{R}_n(k+j), k \leq k_n - j) \\ & \leq C \frac{(1+2d_n)(1+\log(j \wedge ((k_n-j)+1)))^2}{\sqrt{b_n} (k_n-j)^{\frac{3}{2}}}, \end{aligned}$$

therefore, we conclude that

$$\begin{aligned} \mathbb{E}(\mathcal{E}_{n,R_n}(A) - \mathcal{E}_{n,F_n}(A)) & \leq C \frac{(1+2d_n)}{\sqrt{b_n}} k_n^{\frac{3}{2}} \sum_{j \leq k_n} \frac{(1+(\log(j \wedge ((k_n-j)+1))))^4}{((k_n-j)+1)^{\frac{3}{2}} j^{\frac{3}{2}}} \\ & \leq 2C \frac{(1+2d_n)}{\sqrt{b_n}} \sum_{j \leq [k_n/2]} \frac{(1+\log(j))^4}{j^{\frac{3}{2}}} \rightarrow_{n \rightarrow \infty} 0, \end{aligned}$$

as $\sum_{j \leq [k_n/2]} \frac{(1+\log(j))^4}{j^{\frac{3}{2}}} < \infty$. □

This lemma implies that the two extremal processes \mathcal{E}_{n,F_n} and \mathcal{E}_{n,R_n} have the same weak limit. Consequently, the the same as the one of $\mathcal{E}_n^{b_n}$. For $u \in \mathcal{T}^{(n)}$, we introduce

$$H_n(u) = \{S_{u_k} \leq F_n(k), k \leq k_n\}$$

the set of individuals satisfying the F_n -barriers. We will prove that, with high probability, the set of pairs of extremal particles that branch off at time $k \geq 1$ and stay all the time below the barrier $k \mapsto F_n(k)$ is vanishing in the large n -limit. This show that all particles contributing in the extremal process split from the root.

Lemma 16. *With the same notation used before, we have,*

$$\lim_{n \rightarrow \infty} \mathbb{E}(\#\{(u, v), |u \wedge v| \geq 1, H_n(u), H_n(v), S_u - m_n \in A, S_v - m_n \in A\}) = 0.$$

Proof. By considering the positions of any pairs of individuals (u, v) at the generation k_n and at their common ancestors $u \wedge v$ we have

$$\begin{aligned} & \mathbb{E}(\#\{(u, v), |u \wedge v| \geq 1, S_u - m_n \in A, S_v - m_n \in A, H_n(u), H_n(v)\}) \\ & = \mathbb{E} \left(\sum_{j=1}^{k_n-1} \sum_{|w|=j} \mathbf{1}_{\{S_{w_i} \leq F_n(i), i \leq j\}} \sum_{(u_{j+1}, v_{j+1})} \sum_{(u, v)} \mathbf{1}_{\{S_u - m_n \in A, S_v - m_n \in A, S_{u_k} \leq F_n(k), S_{v_k} \leq F_n(k), j+1 \leq k \leq k_n\}} \right) \end{aligned}$$

where the double sum $\sum_{(u_{j+1}, v_{j+1})}$ is over pairs (u_{j+1}, v_{j+1}) of distinct children of $w = u \wedge v$ and $\sum_{(u, v)}$ is over pairs (u, v) such that $|u| = |v| = k_n$ and u is a descendant of u_{j+1} , and v is a descendant of v_{j+1} . Applying the Markov property at time $j+1$ we get

$$\begin{aligned} & \mathbb{E}(\#\{(u, v), |u \wedge v| \geq 1, S_u^{b_n} - m_n \in A, S_v^{b_n} - m_n \in A, H_n(u), H_n(v)\}) \\ & \leq \mathbb{E} \left(\sum_{j=1}^{k_n-1} \sum_{|w|=j} \mathbf{1}_{\{S_{w_i} \leq F_n(i), i \leq j\}} \sum_{(u_{j+1}, v_{j+1})} \mathbf{1}_{\{S_{u_{j+1}} \leq F_n(j+1), S_{v_{j+1}} \leq F_n(j+1)\}} \phi_{j,n}(S_{u_{j+1}}) \phi_{j,n}(S_{v_{j+1}}) \right), \end{aligned} \tag{42}$$

where

$$\phi_{j,n}(z) = \mathbb{E} \left(\sum_{|u|=k_n-j-1} \mathbf{1}_{\{z+S_u - m_n \in A, S_{u_k} + z \leq F_n(j+k+1), k \leq k_n-j-1\}} \right).$$

Now using Proposition 3 we obtain,

$$\phi_{j,n}(z) = \mathbb{E} \left(e^{-\theta^* \bar{T}_{k_n-j-1}} \mathbf{1}_{\{z+\bar{T}_{k_n-j-1}-m_n+(k_n-j-1)b_n v \in A, \bar{T}_k + z \leq F_n(j+k+1)-kb_n v, k \leq k_n-j-1\}} \right).$$

Summing with respect to the value of $\bar{T}_{k_n-j-1} - m_n + (k_n - j - 1)b_nv$ we have

$$\begin{aligned} \phi_{j,n}(z) &\leq \frac{Cn^{\frac{3}{2}}}{b_n} e^{\theta^*(z-(j+1)b_nv)} \sum_{h \geq 0} e^{-\theta^* h} \\ &\times \mathbb{P}_{z-(j+1)b_nv}(\bar{T}_{k_n-j-1} - a_n - a \in [h, h+1], \bar{T}_k \leq F_n(j+k+1) - kb_nv, k \leq k_n - j - 1). \end{aligned} \quad (43)$$

Note that If **(H₂)** holds by Theorem 9, we bound the quantity (43) by

$$\begin{aligned} &\mathbb{P}_z(\bar{T}_{k_n-j-1} - m_n + (k_n - j - 1)b_nv - a \in [h, h+1], \bar{T}_k \leq F_n(j+k+1) - kb_nv, k \leq k_n - j - 1) \\ &\leq \mathbb{P}_{z-(j+1)b_n}(\mathcal{D}_n^{(k)}), \end{aligned}$$

where

$$\mathcal{D}_n^{(k)} = \left\{ \sqrt{b_n} \hat{S}_{k_n-j-1} - a_{k_n-j-1} - a \in [h - d_n, h + 1 + d_n], \sqrt{b_n} \hat{S}_k \leq \bar{F}_n(k) + d_n, k \leq k_n - j - 1 \right\}.$$

Now assume that either **(H₁)** or **(H₂)** hold. Thanks to Lemma 7 or 11 we have,

$$\phi_{j,n}(z) \leq C(1 + 2d_n)k_n^{\frac{3}{2}} e^{\theta^*(z-(j+1)b_nv)} \frac{(1 - \frac{z-(j+1)b_nv}{\sqrt{b_n}})^2}{(k_n - j)^{\frac{3}{2}}}.$$

By replacing this in the equation (42), we get

$$\begin{aligned} &\mathbb{E}(\#((u, v), |u \wedge v| \geq 1, S_u^{b_n} - m_n \in A, S_v^{b_n} - m_n \in A, H_n(u), H_n(v))) \\ &\leq C(1 + 2d_n)k_n^{\frac{3}{2}} \mathbb{E}\left(\sum_{j=1}^{k_n-1} \frac{1}{(k_n - j)^{\frac{3}{2}}} \sum_{|w|=j} \mathbf{1}_{\{S_{w_i} \leq F_n(i), i \leq j\}} \times \right. \\ &\quad \left. \sum_{(u_{j+1}, v_{j+1})} \mathbf{1}_{\{S_{u_{j+1}} \leq F_n(j+1), S_{v_{j+1}} \leq F_n(j+1)\}} e^{\theta^* S_{u_{j+1}} - (j+1)b_nv + S_{v_{j+1}} - (j+1)b_nv} f_{n,j}(S_{u_{j+1}}) f_{n,j}(S_{v_{j+1}})\right) \end{aligned}$$

where $f_{n,j}(u) = (1 + \frac{z-(j+1)b_nv}{\sqrt{b_n}})^2$. In the other hand we can bound the double sum

$$\sum_{(u_{j+1}, v_{j+1})} \mathbf{1}_{\{S_{u_{j+1}} \leq F_n(j+1), S_{v_{j+1}} \leq F_n(j+1)\}} e^{\theta^* [S_{u_{j+1}} - (j+1)b_nv + S_{v_{j+1}} - (j+1)b_nv]} f_{n,j}(S_{u_{j+1}}) f_{n,j}(S_{v_{j+1}})$$

by

$$\begin{aligned} &(1 + \frac{S_w - jb_nv}{\sqrt{b_n}})^4 e^{2\theta^*(S_w - jb_nv)} \\ &\times \mathbb{E}\left(\sum_{\substack{|u|=|v|=1 \\ u \neq v}} (1 + \frac{X_u^{(n)} - b_nv}{\sqrt{b_n}})^2 (1 + \frac{X_v^{(n)} - b_nv}{\sqrt{b_n}})^2 e^{\theta^*(X_u^{(n)} - b_nv + X_v^{(n)} - b_nv)}\right). \end{aligned}$$

Using independence between $X_u^{(n)}$ and $X_v^{(n)}$ for $u \neq v$ and the fact that

$$\mathbb{E}\left(\sum_{|u|=1} (X_u^{(n)} - b_nv) e^{\theta^*(X_u^{(n)} - b_nv)}\right) = 0,$$

we have

$$\begin{aligned} &\mathbb{E}\left(\sum_{\substack{|u|=|v|=1 \\ u \neq v}} (1 + \frac{X_u^{(n)} - b_nv}{\sqrt{b_n}})^2 (1 + \frac{X_v^{(n)} - b_nv}{\sqrt{b_n}})^2 e^{\theta^*(X_u^{(n)} - b_nv + X_v^{(n)} - b_nv)}\right) \\ &\leq \mathbb{E}\left(\sum_{\substack{|u|=|v|=1 \\ u \neq v}} (1 + (\frac{X_u^{(n)} - b_nv}{\sqrt{b_n}})^2) e^{\theta^*(X_u^{(n)} - b_nv)} (1 + (\frac{X_v^{(n)} - b_nv}{\sqrt{b_n}})^2) e^{\theta^*(X_v^{(n)} - b_nv)}\right), \end{aligned}$$

then conditioning on Z_{b_n} and using the following properties

$$\mathbb{E}\left((X_u^{(n)} - b_n v)^2 e^{\theta^*(X_u^{(n)} - b_n v)}\right) = b_n \Lambda''(\theta^*) = b_n \sigma^2 \text{ and } \mathbb{E}\left(e^{\theta^*(X_u^{(n)} - b_n v)}\right) = m^{-b_n},$$

we obtain

$$\begin{aligned} & \mathbb{E}\left(\sum_{\substack{|u|=|v|=1 \\ u \neq v}} \left(1 + \left(\frac{X_u^{(n)} - b_n v}{\sqrt{b_n}}\right)^2\right) e^{\theta^*(X_u - b_n v)} \left(1 + \left(\frac{X_v^{(n)} - b_n v}{\sqrt{b_n}}\right)^2\right) e^{\theta^*(X_v^{(n)} - b_n v)} \middle| Z_{b_n}\right) \\ & \leq (1 + \sigma^2)^2 Z_{b_n} (Z_{b_n} - 1) m^{-2b_n}, \end{aligned}$$

as a consequence, there exists a constant $C > 0$ such that

$$\mathbb{E}\left(\sum_{\substack{|u|=|v|=1 \\ u \neq v}} \left(1 + \left(\frac{X_u^{(n)} - b_n v}{\sqrt{b_n}}\right)^2\right) e^{\theta^*(X_u - b_n v)} \left(1 + \left(\frac{X_v^{(n)} - b_n v}{\sqrt{b_n}}\right)^2\right) e^{\theta^*(X_v - b_n v)}\right) \leq C,$$

which leads to the following inequality

$$\begin{aligned} & \mathbb{E}\left(\#((u, v), |u \wedge v| \geq 1, S_u^{b_n} - m_n \in A, S_v^{b_n} - m_n \in A, H_n(u), H_n(v))\right) \\ & \leq C(1 + 2d_n) k_n^{\frac{3}{2}} \mathbb{E}\left(\sum_{j=1}^{k_n-1} \frac{1}{(k_n - j)^{\frac{3}{2}}} \sum_{|w|=j} \left(1 + \frac{S_w - j b_n v}{\sqrt{b_n}}\right)^4 e^{2\theta^*(S_w - j b_n v)} \mathbf{1}_{\{S_{w_i} \leq F_n(i), i \leq j\}}\right). \end{aligned}$$

On the other hand by Proposition 3 we obtain

$$\begin{aligned} & \mathbb{E}\left(\#((u, v), |u \wedge v| \geq 1, S_u - m_n \in A, S_v - m_n \in A, H_n(u), H_n(v))\right) \\ & \leq C(1 + 2d_n) k_n^{\frac{3}{2}} \mathbb{E}\left(\sum_{j=1}^{k_n-1} \frac{1}{(k_n - j)^{\frac{3}{2}}} \left(1 + \frac{\bar{T}_j}{\sqrt{b_n}}\right)^4 e^{\theta^* \bar{T}_j} \mathbf{1}_{\{\bar{T}_i \leq \bar{F}_n(i), i \leq j\}}\right). \end{aligned}$$

Summing with respect to the values of $\bar{T}_j - \bar{F}_n(j)$, by Lemma 7 or 11 and for n large enough we get

$$\begin{aligned} & \mathbb{E}\left(\#((u, v), |u \wedge v| \geq 1, S_u - m_n \in A, S_v - m_n \in A, H_n(u), H_n(v))\right) \\ & \leq C(1 + 2d_n) k_n^{\frac{3}{2}} \sum_{j=1}^{k_n-1} \frac{1}{(k_n - j)^{\frac{3}{2}}} \sum_{r=0}^{\infty} \mathbb{E}\left(e^{\theta^* \bar{T}_j} \left(1 + \frac{\bar{T}_j}{\sqrt{b_n}}\right)^4 \mathbf{1}_{\{\bar{T}_i \leq \bar{F}_n(i), \bar{T}_j - \bar{F}_n(j) \in [-r-1, -r], i \leq j\}}\right) \\ & \leq C(1 + 2d_n) e^{-\theta^* c_n} k_n^{\frac{3}{2}} \sum_{j=1}^{k_n-1} \frac{1}{(k_n - j)^{\frac{3}{2}}} \sum_{r=0}^{\infty} (1+r)^4 e^{-\theta^* r} \\ & \times \mathbb{P}(\bar{T}_i \leq \bar{F}_n(i), \bar{T}_j - \bar{F}_n(j) \in [-r-1, -r], i \leq j) \\ & \leq \frac{C(1 + 2d_n)}{\sqrt{b_n}} e^{-\theta^* c_n} \sum_{j=1}^{\lfloor \frac{k_n}{2} \rfloor} \frac{k_n^{\frac{3}{2}}}{(k_n - j)^{\frac{3}{2}} j^{\frac{3}{2}}} \rightarrow_{n \rightarrow \infty} 0, \end{aligned}$$

where we used that $\sum_{j=1}^{\lfloor \frac{k_n}{2} \rfloor} \frac{k_n^{\frac{3}{2}}}{(k_n - j)^{\frac{3}{2}} j^{\frac{3}{2}}} \leq 2^{\frac{3}{2}} \sum_{j=1}^{\lfloor \frac{k_n}{2} \rfloor} \frac{1}{j^{\frac{3}{2}}} < \infty$. □

Now we are ready to prove our main result.

Proof of Theorem 1. Let $\phi \in \mathcal{C}_b^{l,+}$, with support $A = [a, \infty)$ where $a \in \mathbb{R}$. We have to show that

$$\lim_{n \rightarrow \infty} \mathbb{E}\left(e^{-\sum_{u \in \mathcal{H}_{k_n}} \phi(S_u^{(n)} - m_n)}\right) = \mathbb{E}\left(\exp\left(-Z_\infty \frac{1}{\sqrt{2\pi\sigma^2}} \int e^{-\theta^* y} (1 - e^{-\phi(y)}) dy\right)\right). \quad (44)$$

First introduce

$$G_n = \{\#(u, v), |u \wedge v| \geq 1, S_u - m_n \geq a, S_v - m_n \geq a\}.$$

By Lemma 14 and Proposition 15, it is enough to prove (44) for the extremal process

$$\mathcal{E}_{n,F_n} = \sum_{u \in \mathcal{H}_{k_n}} \delta_{S_u - m_n} \mathbf{1}_{H_n}.$$

where $H_n = \{S_{u_k} \leq F_n(k), \forall k \leq k_n\}$. Using Lemma 16, we have $\mathbb{P}(G_n^c) \rightarrow_{n \rightarrow \infty} 0$, therefore

$$\begin{aligned} & \mathbb{E} \left(\exp - \left(\sum_{u \in \mathcal{H}_{k_n}} \phi(S_u - m_n) \mathbf{1}_{H_n} \right) \right) \\ &= \mathbb{E} \left(\mathbf{1}_{G_n} \exp - \left(\sum_{u \in \mathcal{H}_{k_n}} \phi(S_{u_{k_n}} - m_n) \mathbf{1}_{H_n} \right) \right) + o(1) \\ &= \mathbb{E} \left(\mathbf{1}_{G_n} \prod_{|w|=1} \left(\exp - \left(\sum_{\substack{u > w \\ |u|=k_n}} \phi(S_{u_{k_n}} - m_n) \mathbf{1}_{H_n} \right) \right) \right) + o(1). \end{aligned}$$

Observe that $\exp(-\sum_{i=1}^n x_i) = 1 + \sum_{i=1}^n (e^{-x_i} - 1)$ if there exists at most one i such that $x_i \neq 0$. Hence using that under G_n , for all w at the first generation, at most one descendant reaches level m_n , we get,

$$\begin{aligned} & \mathbb{E} \left(\exp - \left(\sum_{u \in \mathcal{H}_{k_n}} \phi(S_u - m_n) \mathbf{1}_{H_n} \right) \right) \\ &= \mathbb{E} \left(\mathbf{1}_{G_n} \prod_{|w|=1} \left(1 + \sum_{\substack{u > w \\ |u|=k_n}} (\exp(-\phi(S_u - m_n + S_w)) - 1) \mathbf{1}_{H_n} \right) \right) + o(1) \\ &= \mathbb{E} \left(\prod_{|w|=1} \left(1 + \sum_{\substack{u > w \\ |u|=k_n}} (\exp(-\phi(S_u - m_n + S_w)) - 1) \mathbf{1}_{H_n} \right) \right) + o(1), \end{aligned}$$

using again that G_n^c is an event of asymptotically small probability, and that this product of random variable is bounded by 1. We now apply the Markov property at time one to obtain

$$\mathbb{E} \left(\exp - \left(\sum_{u \in \mathcal{H}_{k_n}} \phi(S_u - m_n) \mathbf{1}_{H_n} \right) \right) = \mathbb{E} \left(\prod_{|w|=1} 1 + \psi_n(S_w) \mathbf{1}_{\{S_w - b_n v \leq \bar{F}_n(1)\}} \right) + o(1), \quad (45)$$

where

$$\psi_n : x \mapsto \mathbb{E} \left(\sum_{|u|=k_n-1} \left(e^{-\phi(S_u - m_n + x)} - 1 \right) \mathbf{1}_{\{S_{u_k} \leq F_n(k+1) - x, k < k_n\}} \right), \quad (46)$$

using Proposition 3 we have,

$$\psi_n(x) = \mathbb{E} \left(e^{-\theta^* \bar{T}_{k_n-1}} e^{-\phi(\bar{T}_{k_n-1} - \alpha_{b_n} + x)} - 1 \right) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k+1) - (x - b_n v), k < k_n\}}$$

where $\alpha_{b_n} = b_n v - \frac{3}{2\theta^*} \ln(n) + \frac{\ln(b_n)}{\theta^*}$.

Using Lemmas 8 or Corollary 13, depending on whatever we work under (\mathbf{H}_1) or (\mathbf{H}_2) , we obtain the following approximation,

$$\begin{aligned} & \mathbb{E} \left(e^{-\theta^* \bar{T}_{k_n-1}} e^{-\phi(\bar{T}_{k_n-1} - \alpha_{b_n} + x)} - 1 \right) \mathbf{1}_{\{\bar{T}_k \leq \bar{F}_n(k+1) - (x - b_n v), k < k_n\}} \\ & \sim_{n \rightarrow \infty} \frac{1}{\sqrt{2\pi\sigma^2}} L\left(\frac{-(x - b_n v)}{\sqrt{b_n}}\right) e^{\theta^*(x - b_n v)} \int e^{-\theta^* y} (e^{-\phi(y)} - 1) dy. \end{aligned}$$

Plugging this in Equation (45) we get

$$\begin{aligned} & \mathbb{E} \left(e^{-\mathcal{E}_{n, F_n}(\phi)} \right) \\ & \sim_{n \rightarrow \infty} \mathbb{E} \left(\prod_{|w|=1} \mathbb{E} \left(1 + L \left(-\frac{S_w - b_n v}{\sqrt{b_n}} \right) e^{\theta^* (S_w - b_n v)} \mathbf{1}_{\{S_w - b_n v \leq \bar{F}_n(1)\}} \frac{1}{\sqrt{2\pi\sigma^2}} \int e^{-\theta^* y} (e^{-\phi(y)} - 1) dy \right) \right). \end{aligned}$$

Recall that $v = \frac{\kappa(\theta^*)}{\theta^*}$, then by Girsanov transform we have

$$\mathbb{E} \left(L \left(\frac{-(S_w - b_n v)}{\sqrt{b_n}} \right) e^{\theta^* (S_w - b_n v)} \mathbf{1}_{\{S_w - b_n v \leq \bar{F}_n(1)\}} \right) = m^{-b_n} \mathbb{E} \left(L \left(\frac{S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq \frac{\bar{F}_n(1)}{\sqrt{b_n}} \right\}} \right)$$

where $S_1^{(n)}$ is a centred random walk with variance b_n . Fix $A > 0$

$$\begin{aligned} & \mathbb{E} \left(L \left(\frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq \bar{F}_n(1) \right\}} \right) \\ & = \mathbb{E} \left(L \left(\frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ -A \leq \frac{S_1^{(n)}}{\sqrt{b_n}} \leq \frac{\bar{F}_n(1)}{\sqrt{b_n}} \right\}} \right) + \mathbb{E} \left(L \left(\frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq -A \right\}} \right). \end{aligned}$$

As the function $x \mapsto L(x) \mathbf{1}_{\{x \in [-A, \frac{\bar{F}_n(1)}{\sqrt{b_n}}]\}}$ is bounded, by central limit theorem we have

$$\lim_{n \rightarrow \infty} \mathbb{E} \left(L \left(\frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ -A \leq \frac{S_1^{(n)}}{\sqrt{b_n}} \leq \frac{\bar{F}_n(1)}{\sqrt{b_n}} \right\}} \right) = \frac{1}{\sqrt{2\pi}} \int_{-A}^0 L(-y) e^{-\frac{y^2}{2}} dy. \quad (47)$$

Additionally, we have by Equation (16) and the Markov inequality

$$\begin{aligned} & \left| \mathbb{E} \left(L \left(\frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq -A \right\}} \right) \right| \leq C \left| \mathbb{E} \left(\left(1 + \frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq -A \right\}} \right) \right| \\ & \leq C \frac{\mathbb{E}(|S_1^{(n)}|)}{\sqrt{b_n} A} + \mathbb{E} \left(\frac{|S_1^{(n)}|}{\sqrt{b_n} A} \frac{|S_1^{(n)}|}{\sqrt{b_n}} \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq -A \right\}} \right) \leq \frac{C}{A}. \end{aligned}$$

Therefore, we have

$$\limsup_{n \rightarrow \infty} \left| \mathbb{E} \left(L \left(\frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq \frac{\bar{F}_n(1)}{\sqrt{b_n}} \right\}} \right) - \frac{1}{\sqrt{2\pi}} \int_{-A}^0 L(-y) e^{-\frac{y^2}{2}} dy \right| \leq \frac{C}{A}.$$

Thus, letting $A \rightarrow \infty$ in (47), by (14) we obtain that

$$\begin{aligned} & \lim_{n \rightarrow \infty} \mathbb{E} \left(L \left(\frac{-S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \leq \frac{\bar{F}_n(1)}{\sqrt{b_n}} \right\}} \right) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^0 L(-y) e^{-\frac{y^2}{2}} dy \\ & = \mathbb{E} \left(L \left(\frac{S_1^{(n)}}{\sqrt{b_n}} \right) \mathbf{1}_{\left\{ \frac{S_1^{(n)}}{\sqrt{b_n}} \geq 0 \right\}} \right) = L(0) = 1. \end{aligned}$$

As a consequence we obtain

$$\begin{aligned} & \mathbb{E} \left(\prod_{|w|=1} \mathbb{E} \left(1 + L \left(-\frac{S_w - b_n v}{\sqrt{b_n}} \right) e^{\theta^* (S_w - b_n v)} \mathbf{1}_{\{S_w - b_n v \leq \frac{\bar{F}_n(1)}{b_n}\}} \frac{1}{\sqrt{2\pi\sigma^2}} \int e^{-\theta^* y} (e^{-\phi(y)} - 1) dy \right) \right) \\ & \sim_{n \rightarrow \infty} \mathbb{E} \left(\left(1 - \frac{1}{\sqrt{2\pi\sigma^2}} m^{-b_n} \int e^{-\theta^* y} (1 - e^{-\phi(y)}) dy \right)^{\#\{u, |u|=1\}} \right) \\ & = \mathbb{E} \left(\left(1 - \frac{1}{\sqrt{2\pi\sigma^2}} m^{-b_n} \int e^{-\theta^* y} (1 - e^{-\phi(y)}) dy \right)^{Z_{b_n}} \right). \end{aligned}$$

Finally, applying dominated convergence theorem and by assumption (3) we deduce that (44) holds for all function $\phi \in \mathcal{C}_b^{l,+}$, which concludes the proof using Remark 2. \square

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