

# Replacements of small- by large-ranged species scale up to diversity loss in Europe's temperate forest biome

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1 Replacements of small- by large-ranged species scale up to diversity loss in Europe's temperate forest biome 2 Ingmar R. Staude<sup>1,2</sup>, Donald M. Waller<sup>3</sup>, 3 Markus Bernhardt-Römermann<sup>4</sup>, Anne D. Bjorkman<sup>5</sup>, Jörg Brunet<sup>6</sup>, Pieter De Frenne<sup>7</sup>, Radim Hédl<sup>8,9</sup>, Ute 4 Jandt<sup>1,2</sup>, Jonathan Lenoir<sup>10</sup>, František Máliš<sup>11,12</sup>, Kris Verheyen<sup>7</sup>, Monika Wulf<sup>13,14</sup>, Henrique M Pereira<sup>1,2</sup>, 5 Pieter Vangansbeke<sup>7</sup>, Adrienne Ortmann-Ajkai<sup>15</sup>, Remigiusz Pielech<sup>16</sup>, Imre Berki<sup>17</sup>, Markéta 6 Chudomelová<sup>8</sup>, Guillaume Decocq<sup>10</sup>, Thomas Dirnböck<sup>18</sup>, Tomasz Durak<sup>19</sup>, Thilo Heinken<sup>14</sup>, Bogdan 7 Jaroszewicz<sup>20</sup>, Martin Kopecký<sup>21,22</sup>, Martin Macek<sup>21</sup>, Marek Malicki<sup>23</sup>, Tobias Naaf<sup>13</sup>, Thomas A. Nagel<sup>24</sup>, 8 Petr Petřík<sup>21</sup>, Kamila Reczyńska<sup>23</sup>, Fride Høistad Schei<sup>25</sup>, Wolfgang Schmidt<sup>26</sup>, Tibor Standovár<sup>27</sup>, Krzysztof 9 Świerkosz<sup>28</sup>, Balázs Teleki<sup>29,30</sup>, Hans Van Calster<sup>31</sup>, Ondřej Vild<sup>8</sup>, Lander Baeten<sup>7</sup> 10 11 Affiliations 12 13 1. German Centre for Integrative Biodiversity Research (iDiv) Halle-Jena-Leipzig, Leipzig, Germany 14 Institute of Biology, Martin-Luther University Halle-Wittenberg, Halle (Saale), Germany 15 3. Department of Botany, University of Wisconsin, Madison, WI, USA 16 4. Institute of Ecology and Evolution, Friedrich Schiller University Jena, Jena, Germany 17 5. Senckenberg Gesellschaft für Naturforschung, Biodiversity and Climate Research Centre (SBiKF), Frankfurt, 18 Germany 19 6. Southern Swedish Forest Research Centre, Swedish University of Agricultural Sciences, Alnarp, Sweden 20 7. Forest & Nature Lab, Ghent University, Gontrode, Belgium 21 8. Institute of Botany, Czech Academy of Sciences, Brno, Czech Republic 9. Department of Botany, Faculty of Science, Palacký University in Olomouc, Olomouc, Czech Republic 22 23 10. EDYSAN - Ecologie et Dynamique des Systèmes Anthropisés, Jules Verne University of Picardie, Amiens, 24 France 25 11. Faculty of Forestry, Technical University in Zvolen, Zvolen, Slovakia 26 12. National Forest Centre, Zvolen, Slovakia 13. Leibniz Centre Centre for Agricultural Landscape Research (ZALF), Müncheberg, Germany 27 28 14. Institute of Biochemistry and Biology, University of Potsdam, Germany 29 15. Department of Hidrobiology, University of Pécs, Szekszárd, Hungary 30 16. Department of Forest Biodiversity, Faculty of Forestry, University of Agriculture in Kraków, Poland 31 17. University of Sopron, Faculty of Forestry, Institute of Environmental and Earth Sciences, Sopron, Hungary 32 18. Environment Agency Austria, Vienna, Austria 33 19. Department of Plant Physiology and Ecology, University of Rzeszów, Rzeszów, Poland 34 20. Białowieża Geobotanical Station, Faculty of Biology, University of Warsaw, Poland 35 21. Institute of Botany of the Czech Academy of Sciences, Průhonice, Czech Republic 36 22. Faculty of Forestry and Wood Sciences, Czech University of Life Sciences Prague, Kamýcká, Czech Republic 37 23. Department of Botany, Faculty of Biological Sciences, University of Wrocław, Poland 38 24. Department of forestry and renewable forest resources, University of Ljubljana, Ljubljana, Slovenia 39 25. Norwegian Institute of Bioeconomy Research, Division of Forest and Forest Resources, Bergen, Norway

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Biodiversity time series reveal global losses and accelerated redistributions of species, yet no net loss in local species richness. To better understand how these patterns are linked, we quantify how individual species trajectories scale up to diversity changes using data from 68 vegetation resurvey studies of seminatural forests in Europe. Herb-layer species with small geographic ranges are being replaced by more widely distributed species and our results suggest this is less due to species abundances than to species nitrogen (N) niches. N-deposition accelerates extinctions of small-ranged, N-efficient plants and colonization by broadly distributed, N-demanding plants including non-natives. Despite no net change in species richness at the spatial scale of a study site, losses of small-ranged species reduce biome-scale (gamma) diversity. These results provide one mechanism to explain the directional replacement of small-ranged species within sites and thus patterns of biodiversity change across spatial scales.

Introduction

The biological diversity on Earth is changing due to human activities. At the global scale, species are going extinct at rates that signal a mass extinction<sup>1,2</sup>. Per contra, at local scales whether or not diversity is declining is controversial. Time-series studies find that sites may gain or lose species with no directional global trend<sup>3–5</sup>. Space-for-time comparisons find substantial losses in local diversity globally due to human land use<sup>6</sup>. While spatial comparisons are criticized for neglecting that community dynamics are much

slower than the speed of environmental changes<sup>7</sup>, time-series studies are challenged for not being spatially representative of human land use effects<sup>5,8,9</sup>. Yet even in relatively intact places and independent from changes in local diversity, species appear to be replacing each other more rapidly than predicted from only natural changes<sup>4,10</sup>. These local-scale replacements alone could lead to species loss at larger spatial scales, if species with small geographic ranges are frequently replaced by species with larger ranges. Tests of this prediction remain, however, sparse. Why should small-ranged species decline relative to those with larger ranges within sites? A greater tendency of species with small ranges to decline or go locally extinct could reflect lower abundance, greater ecological specificity (narrower niches), or both. Species with small geographic ranges generally tend to have smaller local populations<sup>11–13</sup>, and with decreasing population size, vulnerability to environmental change increases<sup>14</sup>. Species with small range size also tend to be more specialized with narrower niche breadth<sup>15</sup> and may therefore lack flexibility to cope with anthropogenic changes in their abiotic and biotic environment. As these changes increase, we might expect niche effects to strengthen leading to high species replacement. For plants, a primary limiting factor in many natural environments is nitrogen (N)<sup>16</sup>. Humans have substantially altered the distribution and availability of N over the last century<sup>17</sup>. Chronic high Ndeposition has now saturated many ecological systems, exceeding critical loads<sup>18-21</sup>. Increases in a key resource like N reorder competitive relationships among plant species within communities, favoring Ndemanding species at sites of high N-deposition across many ecosystems<sup>22–28</sup>. Yet, beyond local-scale community changes, how these shifts link to biodiversity change across spatial scales remains largely unexplored. Here, we explore the role of species range size and N-deposition in driving the systematic shifts in species composition and scale-dependent patterns of diversity changes observed in extensive long-term

vegetation data<sup>27,29,30</sup>. The N-niche of species relates to their range size in that species that use N more

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efficiently tend to have smaller ranges relative to N-demanding species<sup>31</sup>. This may reflect the ability of N-demanding species to grow faster<sup>32,33</sup>. Faster-growing species usually have smaller seeds<sup>34</sup> enabling further dispersal<sup>35</sup>, and more widely dispersed species tend to have wider ranges<sup>36</sup>. Increases in N are thus hypothesized to favor larger-ranged species that grow faster under more fertile conditions, allowing them to become superior competitors, reducing the survivorship of N-efficient, small-ranged species. Patterns of global biodiversity loss and local maintenance of diversity would be consistent with such species replacements, with a few large-ranged species replacing many small-ranged species in a process termed biotic homogenization<sup>37</sup>.

We compiled a database of 68 resurvey studies of herb-layer communities from semi-natural, temperate forests spread across Europe (Fig. 1). We use these data to ask: 1. Have small-ranged forest plant species declined over time? 2. Do any such trends simply reflect their lower abundance (given the range size -- abundance correlation), or does it reflect niche effects that strengthen with N-deposition? 3. Do species replacements under N-deposition evoke a homogenization pattern with small-scale richness remaining constant on average while larger-scale richness declines?

Results

Plant species that went extinct from a study site had smaller range sizes than species that persisted and those that colonized. In contrast, persisting and colonizing species had similar range sizes (Fig. 2a and Supplementary Table 2). If this pattern reflected only lower abundance at the study site, range size should not add power for predicting extinctions once abundance is controlled for. Here, we estimate abundance as occupancy across plots within a site<sup>38,39</sup>. Occupancy was a strong predictor of probability of extinction (slope:  $\beta = -3.63$ , standard deviation:  $\sigma = 0.28$ ). Over an average time interval of 38 years, species of average occupancy had a 10% chance of going extinct from a study site relative to up to 60% for species of low occupancy (Fig. 2b and Supplementary Table 3). Range size still had a negative effect on species'

extinction probability even after controlling for occupancy ( $\beta = -0.21$ ,  $\sigma = 0.05$ ). For species of average occupancy, extinction probabilities declined by more than 50% as range size increased (Fig. 2c and Supplementary Table 3). The total effect of range size became only slightly stronger when occupancy was not controlled for  $(\beta = -0.28, \sigma = 0.06)$ , suggesting that only 25 % of the range size effect reflects occupancy (Supplementary Table 3). As species' range size presents a basic summary of the ecological characteristics of species (mainly in terms of climatic and edaphic niches<sup>40,41</sup>), the remaining effects of range size likely reflect aspects of species niches. Higher cumulative N-deposition (ΔN, see Methods for details) between surveys sharply increased probabilities of extinction from a site ( $\beta = 0.37$ ,  $\sigma = 0.15$ ; Fig. 3a, Supplementary Table 4), having accounted for confounding variables such as inter-census time period, study area and latitude (see Methods for all variables). This increase in extinction probability disproportionally affected small-ranged species, as shown by the negative interaction between range size and N-deposition ( $\beta = -0.1$ ,  $\sigma = 0.03$ ). Extinction probability of the species with the smallest range sizes increased from ~4% to ~27% as N-deposition increased from 45 to 721 kg ha<sup>-1</sup> (with other predictors at their mean). In contrast, risks of extinction for large-ranged species were much lower and more stable (Fig. 3b and Supplementary Table 4). These results show that probabilities of extinction increase with N-deposition, mostly due to the loss of small-ranged species. Has this eroded study-level species richness or have increases in colonization balanced these extinctions? We found no systematic shifts in species richness within study sites as Ndeposition increased, again accounting for covariates ( $\beta=0.11$ ,  $\sigma=0.15$ ; Fig. 3c, Extended Data 1 and Supplementary Table 5). This implies that higher levels of N-deposition have facilitated the replacement of small-ranged species. Colonizing species had larger ranges (Fig. 2a) and included several non-native species ( $\beta = 0.44$ ,  $\sigma = 0.16$ ; Fig 3d, Supplementary Table 5). As more N-efficient species went extinct with higher N-deposition ( $\beta=-0.25$ ,  $\sigma=0.12$ ; Fig 3e) and colonizing species generally had higher N-

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demands (Extended Data 2), community composition has shifted systematically towards more N-demanding species ( $\beta=0.28$ ,  $\sigma=0.12$ ; Fig. 3e and Supplementary Table 6).

Although small-scale diversity has not declined, N-deposition may threaten other forms of diversity. As small-ranged, N-efficient species are extirpated and replaced by wide-ranging, non-native and N-demanding species, these forest plant communities have converged in composition. Gilliam<sup>19</sup> predicted that such declines in beta (and thus gamma) diversity would occur as N-deposition tends to increase the spatial homogeneity of nutrient availability. Variation in nutrient availability among sites has indeed declined since the baseline surveys (difference between variances:  $\delta = -0.16$ ,  $\sigma = 0.08$ ) (Fig. 3f and Supplementary Table 7, see Methods for estimation of nutrient availability). This homogenization in nutrient availability appears linked to declines in overall (biome-scale) species richness as gamma-diversity across these 68 sites declined by 4% (from 1,012 to 972 species).

### Discussion

Using large-scale temporal vegetation change datasets, we provide evidence that the geographic range size of species predicts long-term shifts in forest-floor plant communities. Small-ranged species are replaced by those with larger ranges and our results suggest this is more due to species niches than abundances. The loss of small-ranged species amplified under high N-deposition and, consistent with our expectation that species' range size and N-demand positively correlate, communities shifted towards species with higher N-demand. Despite the loss of small-ranged species, the number of species within study sites has not declined in response to increasing N-deposition, suggesting that species losses have been balanced by species gains. Nevertheless, the floristic distinctiveness of these forests erodes as more cosmopolitan and non-native species replace a set of more finely-adapted species. These replacements ran in parallel with the abiotic homogenization resulting from chronic N-deposition and scaled to a loss of biodiversity in Europe's temperate forests in recent decades.

Our study confirms that small population size is a strong predictor of extinction from a site 14,42,43. Yet this did not provide much explanation for the greater extinction risk of small-ranged species as would be expected given a positive range size—abundance relationship. This suggests not all small-ranged plant species in these forests have small populations. In fact, plants show many exceptions to this relationship. For example, 87% of small-ranged species from the British Isles are locally common<sup>44</sup>; conversely one of the largest ranging woody species of the globe, Juniperus communis, has small populations in many regions<sup>45</sup>. Indeed, several plant studies find that range size and abundance do not necessarily covary<sup>45,46</sup>. This suggests that range size affects species' extinction probability mostly via niche rather than demographic effects, a conjecture supported by the fact that N-deposition mostly affected small-ranged (narrow niche) species. Despite declines in small-ranged species, forest sites in our study did not systematically decrease in species richness. This suggests species losses are offset by species gains. This finding echoes other resurvey studies that document little directional temporal trend in small-scale species richness despite increased species turnover<sup>4,10</sup>. But similarly, this finding is likely to not reflect the full impact of intensive human land use<sup>8,9</sup> as our study sites are confined to semi-natural forests. Given that the effect of range size reflects species' niches, species turnover accelerated under N-deposition and communities not only shifted towards larger range size but also towards higher average N-demands. Other studies from forest ecosystems report the same shifts towards more N-demanding species from eutrophication and similarly limited effects of N-deposition on forest-floor plant species richness<sup>24,47–49</sup>. This contrasts with opencanopy ecosystems that, being not primarily light limited, consistently lose species from N-addition through shading<sup>50,51</sup>. Beyond a threshold, however, N-deposition also reduces plant diversity in forests, as shown in North America where roughly a quarter of 14,000+ sites showed susceptibility to N-driven species losses<sup>30</sup>.

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Although we find no evidence of a directional change in species richness within studies, the total number of species across studies (gamma diversity) has declined. The observed 4% decline in gamma-diversity likely underestimates the true species loss in the European temperate forest biome, as studies in our database are not completely spatially representative of key human pressures in Europe. Resurveys have been selected to be especially located in large, historically continuously forested (ancient) and seminatural forests where no land-use change took place prior to the baseline survey and no large change in management occurred between the surveys. Change in land use and other disturbances that open up forest canopies and increase light availability may exacerbate effects of global warming by eliminating the thermal insulating layer that protects understories from thermophilisation<sup>52,53</sup>. However, changes in light availability usually occur at the plot-level as opposed to N-deposition that acts at the scale of an entire study area. Because we evaluated community dynamics at the study-level, we expect that light availability changes do not confound the N-signal we found. Nitrogen releases to the environment remain high in Europe (https://www.eea.europa.eu/airs/2018). Despite recent declines, these emissions still exceed critical loads for most of Europe's protected habitats<sup>54</sup>. It is important to learn what long-lasting effects N-deposition may have on Europe's ecosystems and how reversible these are as emissions decline. In contrast to the rapid recovery of plotscale experimental communities, where species are still present in the area and able to colonize once N additions are ceased<sup>22,55</sup> (but see ref.<sup>56</sup>), the loss of small-ranged species from entire regions is far less

likely to be reversed in the short term. At this point, it is thus unclear whether the declines in N emissions

mandated under the EU National Emission Ceilings Directive (2016/2284/EU) will be feasible or sufficient

to allow the recovery of Europe's plant species and communities.

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Methods

208 Database

We compiled a database containing data from 68 understory resurvey studies distributed across the European temperate deciduous forest biome (see www.forestreplot.ugent.be and ref. <sup>57</sup> for inclusion criteria). These studies cover 15 European countries, from Norway in the north to Slovenia in the south and from Ireland in the west to Poland in the east (Fig. 1). All surveys were conducted in historically continuously forested 'ancient' and semi-natural deciduous forest (sensu ref. <sup>58</sup>). These forests did thus not experience historical land-use change prior to the first (baseline) survey and between the surveys. Furthermore, the forests are mostly managed, but did not experience stand-replacing disturbances between the baseline survey and the most recent resurvey. In each survey, the herbaceous understory was recorded in multiple permanent or quasi-permanent plots. Plot size ranged between 1m² to 1000m² across studies (median: 400 m²). Number of plots ranged from 10 to 190 across studies (median: 43 plots, Supplementary Figure 1b). Plots were allocated across areas ranging from 1ha to c. 2.5x10<sup>6</sup> ha (median: 1700 ha, Supplementary Figure 1d). Baseline surveys were carried out between 1933 and 1999. The most recent resurveys were made between 1969 and 2017 generating time intervals between surveys from 15 to 78 years (median: 42 years, Supplementary Figure 1c). We accounted for changes in taxonomy between surveys by determining the accepted species name for each species using GBIFs backbone taxonomy (gbif.org). Harmonization thus ensured no double-counting species due to synonymy. Our database contains 1,162 species in total.

Species level variables

Species' trajectory

We determined the trajectory of species at the study level. We classified species present in the baseline survey but absent in the resurvey as extinct. Those present in both surveys were classified as persisting. Those absent in the baseline survey but present in the resurvey were classified as colonizing. Resurveys of permanent plots always miss some species, generating pseudo-colonizations and extinctions that can inflate these estimates for rare species<sup>59,60</sup>. We did not correct our estimates of colonization and extinction for bias proposed by Beck<sup>59</sup> as we adjust for initial abundance (occupancy across plots) in our model which is strongly correlated with any such bias<sup>61</sup>. This means our estimates of the effects of initial abundance on extinction may be slightly inflated.

Occupancy

For each study, we calculated the initial occupancy of species. This is the number of plots that a species has occupied in the baseline survey, divided by the total number of plots in that survey. Occupancy approximates abundance because, empirically and for any biologically relevant point process pattern, they strongly and positively correlate at local to regional scales<sup>38,39</sup>.

Range size

Species range sizes were estimated from the species point occurrence records in GBIF (gbif.org, 18 January 2019). In total, c. 100 million geographically referenced records were available for 1,147 species in our database (1.3 % missing species) after excluding unlikely and impossible coordinates<sup>63</sup>. Records were aggregated to a hexagonal grid (ISEA3H) at three spatial grains: 3.6 km², 10.7 km² and 32 km². The number of cells any given species occupies on such grid represents its range size. Range size therefore measures species' area of occupancy (AOO, expressed in km²). Results in the main text are based on range sizes estimated at mid-resolution (10.7 km²). At this resolution, the smallest ranging species was *Poa pannonica* A.Kern. with an estimated AOO of 21.4 km², the species with the largest range was the annual meadow grass, *Poa annua* L., with an AOO of *c*. 1.6x10<sup>6</sup> km² (Extended Data 3).

For our analyses we use AOO and not the extent of occurrence (EOO, which includes also discontinuities in occupancy) as AOO is a markedly better predictor of mean site abundance and population size<sup>13,64</sup>. While AOO measured from GBIF point occurrence data is increasingly used in the scientific literature, incomplete spatial coverage of digital biodiversity data can lead to an underestimation of range sizes<sup>65</sup>. Specifically, Middle and Northern Asia are some of the most data deficient regions of the world<sup>66</sup>. This may be problematic for European species that extend into these regions. We therefore tested how well our estimates of AOO match estimates of EOO for species where range maps from two renowned distribution atlases were available<sup>67,68</sup> (available for 796 species, 31% missing). The distribution ranges were digitized from scanned atlas pages and rasterized on a 20 km x 20 km grid in the Lambert azimuthal equal-area projection to calculate EOO for all 796 species. Spearman's correlation between AOO and EOO was high ( $\rho = 0.71$ ) for these species. As an overall positive correlation might obscure a weak correlation for continental species that extend into data deficient regions, we also tested for range attribution. We identified 155 species as continental using species indicator values for continentality from ref.<sup>69</sup> (species with values >=6 were classified as continental). Excluding those species made the correlation between AOO and EOO only slightly

stronger ( $\rho = 0.74$ ; Extended Data 3). We therefore assume that data limitations are a less material problem for our set of species.

Study-level variables

Nitrogen-deposition

We quantified N-deposition using the EMEP database (https://emep.int/mscw/mscw\_moddata.html) with c.~11 km (0.1°) grid resolution. Here we chose to focus singly on NO<sub>x</sub> deposition for three reasons. 1) Increasing evidence suggest that the two forms of N-deposition, oxidized (NO<sub>x</sub>) and reduced (NH<sub>y</sub>) N-deposition, have differential, habitat-specific effects on plant communities  $^{70-72}$ . Whereas NH<sub>y</sub> is the most important driver for the decline in plant diversity in grasslands, forest vegetation is found to be most responsive to NO<sub>x</sub>  $^{71}$ . 2) Model estimates of NO<sub>x</sub> deposition also have a lower degree of uncertainty and bias than estimates of NH<sub>y</sub>  $^{73}$ . Local-scale variability of NH<sub>y</sub> deposition is considerably higher as most of it is deposited near the source  $^{74}$ , this variability is likely to be poorly reflected when studying N-effects over larger regions as done in this study. 3) NO<sub>x</sub> is spatially correlated with NH<sub>y</sub> deposition in Europe ( $\rho$  =.69 in this study) and thus representative for broader N-effects. We quantified the cumulative wet and dry deposition of oxidized N (hereinafter and in the main text referred to simply as N-deposition) based on the methods described in ref.  $^{49}$ . First, we calculated N-deposition between 1900 and the year of the baseline survey (N<sub>11</sub>), second we quantified the cumulative N-deposition between 1900 and the resurvey (N<sub>12</sub>), and third we calculated the difference, N<sub>12</sub> – N<sub>11</sub>, to quantify N-deposition between surveys (inter-census N-deposition or  $\Delta$ N).  $\Delta$ N ranged from 45 to 721 kg ha  $^{11}$  (Supplementary Figure 1). Results in the main text are based on  $\Delta$ N.

Change in species numbers

Changes in the number of species found in the resurvey versus the baseline survey were quantified as the difference in size of the recorded species pool for each study (Extended Data 1).

282 Change in non-native species

For each study, species were classified as native or non-native. This classification is based on the Global Register of Introduced and Invasive Species (GRIIS; http://www.griis.org). GRIIS lists species that are non-native in a given country. With these species lists we could flag, for each study, species that are non-native in the country of the study. To calculate the change in non-native species between surveys, we subtracted the proportion of non-native species in the total pool of species recorded in the baseline from the proportion of non-native species in the total pool of species recorded in the resurvey. Thus, we quantify the change in percentage points (Supplementary Figure 3b). Calculation of relative change was not possible because frequently there were no non-native species in the baseline survey (26% of studies, Supplementary Figure 3a). The mean number of non-native species in the baseline survey and the resurvey was 5 and 7 (rounded to the next integer), respectively. The mean number of native species was 110 (baseline survey) and 102 (resurvey).

Change in nitrophilous species and nutrient availability

We estimated changes in species' N-demands using Ellenberg's indicator values (EIVs). EIVs were developed for Central Europe and classify species' habitat niches and their peak occurrence along environmental gradients<sup>75</sup>. In particular, we used EIVs for N or more general productivity<sup>76</sup> that classify species growing on the poorest soils (N-number = 1) to species growing on the most productive soils with excessive nutrient availability (N-number= 9) (Extended Data 2). For each study and survey, we averaged N-numbers across species. Because EIVs equally reflect environmental conditions<sup>47</sup>, these average values approximate both the mean N-demand of a community and the nutrient availability at each survey. To quantify the change in a community's mean N-demand, we subtracted the mean N-demand of the baseline community from the mean N-demand of the resurvey community. To better understand what drives changes in communities mean N-demands, we calculated the average N-demand of extinct and colonizing species (Extended Data 2) for each study.

Data analyses

The entire statistical analysis and R-code is provided in the supplementary information as an R markdown file. The rethinking package<sup>77</sup> was used to compile the following models to Stan Hamiltonian Monte Carlo code. For brevity, all models are presented without priors (complete models and R code is available on figshare [https://figshare.com/s/45d71eb77c23c11bc857]).

We first compared range sizes between extinct, persisting and colonizing species, where species' trajectories are defined at the study scale. Clearly range size is non-normal distributed and starkly right skewed. Since normality of the outcome conditional on the covariates is the central assumption of Gaussian linear models and range size was too skewed for Poisson regression, we normalized range size using an order-quantile transformation <sup>78</sup>. We predicted range size  $(r_i)$  with trajectory  $(\beta_{status[i]})$  and allowed each coefficient to vary by each study  $(\alpha_{study[i],status[i]})$ . The mathematical form of the resulting model is:

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$$r_i \sim Normal(\mu_i, \sigma)$$

$$\mu_i = \beta_{status[i]} + \alpha_{study[i], status[i]}$$

We calculated pairwise contrasts (e.g.,  $\beta_{colonizing} - \beta_{extinct}$ ) to compare range sizes between species trajectories.

Next, we asked can species' extinction probability be predicted by species' occupancy at the time of the baseline survey? Here, we only analyzed species present at the baseline survey, omitting colonizing species. The outcome is a 0/1 (Bernoulli) indicator that a given species persisted or became extinct in a study ( $e_i$ ). As the outcome is binomially distributed, we used logistic regression to predict species' extinction probabilities as a function of occupancy. We allowed intercepts to vary with study ID ( $\alpha_{study[i]}$ ) and species ( $\gamma_{species[i]}$ ) and the effect of occupancy ( $f_i$ ) to vary by study ID ( $\beta_{f,study[i]}$ ). The mathematical form of the model is:

$$e_{i} \sim Binomial(1, p_{i})$$

$$logit(p_{i}) = \overline{\alpha} + \alpha_{study[i]} + \gamma_{species[i]} + \beta_{f,study[i]} * f_{i}$$

Controlling for the effect of occupancy, we then added the range size predictor to the previous model. We used the Widely Applicable Information Criterion (WAIC) to decide whether the model's out of sample predictions improve when varying slopes on range size are included<sup>77</sup>. The resulting model is:

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$$e_i \sim Binomial(1, p_i)$$
 
$$logit(p_i) = \overline{\alpha} + \alpha_{study[i]} + \gamma_{species[i]} + \beta_{f,study[i]} * f_i + \beta_r * r_i$$

329 where  $\beta_r$  is the effect of range size  $(r_i)$ .

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We also modelled the effect of range size alone to estimate by how much it decreases occupancy when is part of the model:

$$e_{i} \sim Binomial(1, p_{i})$$
 
$$logit(p_{i}) = \overline{\alpha} + \alpha_{study[i]} + \gamma_{species[i]} + \beta_{r} * r_{i}$$

Next, we sought to explain variation in average extinction probability across species between studies. In particular, we tested the effect of inter-census N-deposition on the outcome. The effect of N-deposition could, however, be confounded by the time between surveys ( $\Delta t$ ). Cumulative N-deposition is a function of  $\Delta t$ , and  $\Delta t$  itself is likely to affect average extinction probability. We therefore include  $\Delta t$  in the model. Furthermore, the year of the baseline survey (t<sub>1</sub>) can influence the outcome. For example, baseline surveys from later years are likely to be associated to higher cumulative N-deposition than those from earlier years. This might have already affected community composition to the extent that fewer extinctions occur in communities that were sampled in later years as these have already lost species. In our data  $t_1$  is strongly and negatively correlated to  $\Delta t$ ; the earlier the year of the baseline survey, the longer the time between surveys (ho=-0.91, Extended Data 4). This correlation is also reflected in a strong negative correlation between cumulative N-deposition at the year of the baseline survey and time between surveys (ho = -0.67; Extended Data 4). A directed acyclic graph of presumed causal links between predictors and response (Extended Data 5) shows that including time between surveys in the model controls for the confounding effect of year of the baseline survey and importantly, for environmental changes that preceded this survey, such as the cumulative N-deposition at the year of the baseline survey. In addition to these potential confounding variables, the number of plots, their size and the size of the area in which surveys were carried out may directly affect the outcome. For instance, a species with occupancy 0.1 occupied 1 or 10 plots in studies of 10 or 100 plots, respectively and demographic fluctuations should be higher in smaller plots/areas that naturally comprise fewer individuals. As this may clearly affect the average extinction probability across species, we included these variables in the model. Finally, we also included latitude as a covariate in order to

account for latitudinal patterns that might be associated to climate change. Together, this generates the followingmodel:

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$$e_{i} \sim Binomial(1, p_{i})$$

$$logit(p_{i}) = \overline{\alpha} + \alpha_{study[i]} + \gamma_{species[i]}$$

$$+ \beta_{f,study[i]} * f_{i} + \beta_{r} * r_{i}$$

$$+ \beta_{n} * n_{i} + \beta_{t} * t_{i} + \beta_{a} * q_{i} + \beta_{s} * s_{i} + \beta_{a} * a_{i} + \beta_{l} * l_{i}$$

- where  $\beta_n$ ,  $\beta_t$ ,  $\beta_q$ ,  $\beta_s$ ,  $\beta_a$ ,  $\beta_l$  are the effects of inter-census N-deposition  $(n_i)$ , inter-census time period  $(t_i)$ , plot size  $(s_i)$ , site area  $(a_i)$  and latitude  $(l_i)$ , respectively.
- We then asked whether any increase in average extinction probability across species due to N-deposition is driven by
  an increasing extinction probability among small-ranged species or simply a generally higher extinction probability
  across all range sizes. For this, we included the interaction effect between N-deposition and range size in the model:

$$e_{i} \sim Binomial(1, p_{i})$$

$$logit(p_{i}) = \overline{\alpha} + \alpha_{study[i]} + \gamma_{species[i]}$$

$$+ \beta_{f,study[i]} * f_{i} + \beta_{r} * r_{i}$$

$$+ \beta_{n} * n_{i} + \beta_{t} * t_{i} + \beta_{q} * q_{i} + \beta_{s} * s_{i} + \beta_{a} * a_{i} + \beta_{l} * l_{i}$$

$$+ \beta_{nr} * n_{i} * r_{i}$$

- 361 where  $\beta_{nr}$  is the slope of the interaction between inter-census N-deposition  $(n_i)$  and range size  $(r_i)$ .
- Until now, we only analyzed the dynamics of species present at the baseline. To these, we added further models to assess effects on colonizing species. We first asked: are changes in species number ( $d_i$ ) explained by N-deposition? Again we controlled for inter-census time period, plot number, plot size, site area and latitude, generating the following model:

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$$d_i \sim Normal(\mu_i, \sigma)$$

$$\mu_i = \alpha + \beta_n * n_i + \beta_t * t_i + \beta_a * q_i + \beta_s * s_i + \beta_a * a_i + \beta_l * l_i$$

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To visualize the effect of N-deposition, we used a predictor residual plot. In these, the outcome is regressed against the variation of N-deposition that is left unexplained by the other predictor variables in the model. Predictor residual plots allow us to display the actual data while controlling for all other predictors. Because the unit of observation in

this model is the study, we have 68 observations. To display the influence of each data point on posterior predictions, we scaled point sizes by their Pareto k value<sup>77</sup>. We then predicted the percentage point change in nonnative species using the same predictors as in the previous model, again using a predictor residual plot to display the results.

We also tested whether community composition shifts towards more N-demanding species with higher N-deposition. For this, we regressed 1) the average N-demand of extinct species  $(m_i)$  and 2) the change in mean N-demand of the entire community  $(w_i)$  against N-deposition:

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$$m_i/w_i \sim Normal(\mu_i, \sigma)$$

$$\mu_i = \alpha + \beta_n * n_i$$

Finally, we tested whether the variance of nutrient availability across studies was greater in the period of the resurveys than in the period of the baseline surveys. Here, the model is:

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$$a_{ij} \sim Normal(\mu_{ij}, \sigma_{ij})$$

$$\mu_{ij} = \beta_0 + \beta_1 * t_{ij}$$

$$\sigma_{ij} = \gamma_0 + \gamma_1 * t_{ij}$$

where  $a_{ij}$  is the availability of N for the ith study and jth survey period,  $\beta_0$  and  $\gamma_0$  are the mean and standard deviation of the baseline survey respectively,  $\beta_1$  and  $\gamma_1$  are the expected difference between mean and standard deviation of the resurvey and baseline survey respectively, and  $t_{ij}$  is a 0/1 indicator for survey period.

**Author Contributions** 

I.R.S., D.M.W. and L.B. conceived the study, with input from the sREplot working group (M.B.R., A.D.B., J.B., P.D.F, R.H., U.J., J.L., F.M., K.V., and M.W.). I.R.S. performed the analyses, with input from D.M.W. and L.B. I.R.S., D.M.W. and L.B. wrote the manuscript, with input and contributions from M.B.R., A.D.B, J.B., P.D.F., R.H., U.J., J.L., F.M., K.V., M.W., H.M.P., P.V., A.O.A., R.M., I.B., M.C., G.D., T. Dirnböck, T. Durak., W.S., T.H., F.H.S., B.J., M.K., M. Macek, M. Malicki, T.N., T.A.N., P.P., K.R., T.S., K.S., B.T., H.V.C., O.V. Authorship order was determined as follows: (1) core

391 authors; (2) sREplot participants (alphabetical) and other major contributors; (3) authors contributing community 392 composition data and to an advanced version of the manuscript (alphabetical). 393 394 Data availability 395 Community change and environmental site-level data are available on figshare, 396 [https://figshare.com/s/45d71eb77c23c11bc857]. Species composition data are available from forestreplot.ugent.be 397 but restrictions apply to the availability of these data, which were used under license for the current study, and so 398 are not publicly available. Data are however available from the authors upon request and with permission of the 399 forestREplot consortium. 400 401 Code availability 402 R code for all analyses is available on figshare [https://figshare.com/s/45d71eb77c23c11bc857]. 403 404 Competing interests 405 The authors declare no competing interests. 406 407 Acknowledgments 408 This paper is an outcome of the sREplot working group supported by sDiv, the Synthesis Centre of the German 409 Centre for Integrative Biodiversity Research (iDiv) Halle-Jena-Leipzig (DFG FZT 118). P.D.F. and P.V. received funding 410 from the European Research Council (ERC) under the European Union's Horizon 2020 research and innovation 411 programme (ERC Starting Grant FORMICA 757833). K.V. received funding through ERC Consolidator Grant 412 PASTFORWARD 614839. M.K. and M.M. were supported by the Czech Academy of Sciences (RVO 67985939). F.M.

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- 581 Figure legends

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- Figure 1: Map of all 68 resurvey studies included in the forestREplot database, the temperate deciduous
- forest biome in Europe (shaded area)<sup>79</sup> and forest cover for the year 2000 (in green)<sup>80</sup>. Light to dark
- shades of green represent forest cover ranging from 0 to 100% at a spatial resolution of 30 meters.

Figure 2: Species that go extinct from a study site have smaller ranges than persisting and colonizing ones.

Even after controlling for site occupancy, species' range size predicts probability of extinction. a, Expected differences in normalized range size between colonizing, persisting and extinct species. b, Effect of species' site occupancy at the time of the baseline survey on probability of extinction, x axis is standardized, so that zero represents the average site occupancy. c, Effect of species' range size on probability of extinction, after controlling for site occupancy (line shows expectation for site occupancy at its mean). Line segments in a represent ±2 standard deviations from the mean. Transparent ribbons in b and c represent the 89% credible intervals for model mean predictions. Model parameters are shown in Supplementary Table 2 and 3.

Figure 3: Small-ranged species drive the increase in average extinction risk from high N-deposition. Although colonizing species sustain species number, composition shifts towards more non-native and N-demanding species. Spatial heterogeneity of nutrient availability decreases over time. a, Effect of inter-census nitrogen deposition ( $\Delta N$ ) on average probability of extinction across species. b, Triptych plot for the effect of range size on probability of extinction at different levels of  $\Delta N$ , holding all other predictors at their mean.  $\Delta N$  levels are minimum (left), mean (center), and maximum (right)  $\Delta N$ . c-d, Predictor residual response plot of the relationship between inter-census nitrogen deposition and change in study-level species richness (c) and percentage point change in non-native species (d). e, Ellenberg indicator values for nitrogen ( $eiv_N$ ) averaged across extinct species and the change in  $eiv_N$  averaged across all species regressed against intercensus nitrogen deposition. f, Violin plot (density curve and boxplot) of nutrient availability (estimated with community mean  $eiv_N$ ) at the time of the baseline survey ( $t_1$ ) and resurvey ( $t_2$ ). Transparent ribbons in a - e represent the 89% credible intervals for model mean predictions. Point size in c - e is scaled by relative LOOIS Pareto k values. Larger points are more influential. Model parameters are shown in Supplementary Table 4, 5, 6 and 7.





