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Survival probability and size of lineages in antibody affinity maturation

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Affinity Maturation (AM) is the process through which our immune system is able to develop potent antibodies against new pathogens it encounters, and is at the base of the efficacy of vaccines. At its core AM is analogous to a Darwinian evolutionary process, where B-cells mutate and are selected on the base of their affinity for an Antigen (Ag), and Ag availability tunes the selective pressure. When this selective pressure is high the number of B-cells quickly decreases and the population risks extinction in what is known as a *population bottleneck*. Here we study the probability for a B-cell lineage to survive this bottleneck as a function of the progenitor affinity for the Ag. Using recursive relations and probability generating functions we derive expressions for the average extinction time and progeny size for lineages that go extinct. We then extend our results to the full population, both in the absence and presence of competition for T-cell help, and quantify the population survival probability as a function of Ag concentration and population initial size.

I. INTRODUCTION

Affinity Maturation (AM) is a biological process through which our Immune System (IS) generates potent Antibodies (Ab) against newly-encountered pathogens. AM is also at the base of the efficacy of vaccination, in which this process is artificially elicited through the administration of a dose of Ag. The biological mechanisms that govern AM are many and complex, and are the object of many excellent reviews [1–8]. Simply speaking, AM works by subjecting a population of B-lymphocytes (or B-cells) to iterative cycles of mutations and selection for Ag binding, which generate a Darwinian evolutionary process that progressively increases their affinity for the Ag. Maturation takes place in Germinal Centers (GCs), microanatomical structures that appear inside of secondary lymphoid organs. They are divided in two areas: the GC Dark Zone (DZ) in which cells divide and mutate,¹ and the Light Zone (LZ) in which they undergo selection. Cells iteratively migrate between these two compartments. Selection in the LZ is completed in two steps. In the first step cells try to bind the Ag, exposed on the surface of Follicular Dendritic Cells (FDCs). In the second step they compete to receive a survival signal from T-follicular helper (Tfh) cells, in the absence of which they undergo apoptosis. Tfh cells are able to probe the amount of Ag captured by B-cells, preferentially delivering this survival signal to the cells that were most successful in capturing Ag. Cells that receive the signal either migrate back to the DZ for additional rounds of mutation and selection, or they can differentiate in Plasma or Memory Cells (PCs/MCs). The former

are responsible for the production of Abs to fight the infection, while the latter confer long-lasting protection by remaining quiescent until the same Ag is encountered again, in which case they reactivate and produce Abs or enter GCs for further maturation.

In spite of the many recent experimental advancements in the study of AM, several open questions still remain to be answered, which have important implications in vaccine design. For example understanding the role of Ag availability in controlling maturation might lead to optimization of Ag dosage in vaccines [12–14]. Given the complexity of this process, computational models represent an invaluable tool to guide our understanding of AM [15, 16]. In this paper we introduce a stochastic model of AM to study the survival probability of B-cell lineages in GCs. Experimental analysis of vaccine-responsive lineages shows signatures of selection in their reconstructed phylogenies [17]. This selection pressure, which is partially controlled by Ag availability [13], is important to push lineages towards maturation, but at the same time an excessive pressure might be deleterious. Indeed, several maturation models present a phenomenology termed *population bottleneck* [18–20], in which strong selection pressure causes a decrease in GC population size, potentially leading to extinction. As a consequence of this trade-off optimal maturation is achieved at intermediate levels of selection pressure. In these models however the survival probability is only quantified using numerical simulations. Here we perform instead a more theoretical analysis. We start by considering the dependence of a lineage survival probability on the progenitor affinity. Through the use of recursive relations and probability generating functions we are able to evaluate this probability, and also quantify extinction time and progeny size for lineages that go extinct. We then extend our approach to analyze the extinction probability for the full B-cell population, and its dependence on Ag concentration and initial population size. Our analysis provides theoretical insight on what controls the survival proba-

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¹ In the DZ B-cells express high levels of *Activation-Induced cytidine Deaminase*, an enzyme that increases the natural rate of DNA mutations up to 10^{-3} per base-pair per generation [9–11].

bility in maturation.

II. MODEL FOR STOCHASTIC MATURATION

Our model for stochastic maturation is inspired from previous works [13, 19]. The model is simple enough to be analytically tractable, while retaining the main aspects of the bottleneck phenomenology.

A. Steps in affinity maturation

We consider the evolution of a population of B-cells inside a GC. Through repeated cycles of mutation and selection the population increases its average affinity for the Ag over time. In our model each cell in the population is solely characterized by its affinity for the Ag, measured in terms of binding energy ϵ and expressed in units of $k_B T$.

The simulation starts when the GC is mature (roughly 1 week after Ag injection [2]). The initial population is composed of N_i cells whose binding energy is independently extracted from a Gaussian distribution of naive responders, with mean μ_i and standard deviation σ_i . Cells undergo iterative rounds of duplication, mutation and selection. These steps are schematized in fig. 1.

At the beginning of the round cells duplicate once in the GC DZ. Each daughter cell can then independently

- undergo an affinity-affecting mutation with probability p_{aa} , which causes its binding energy to change

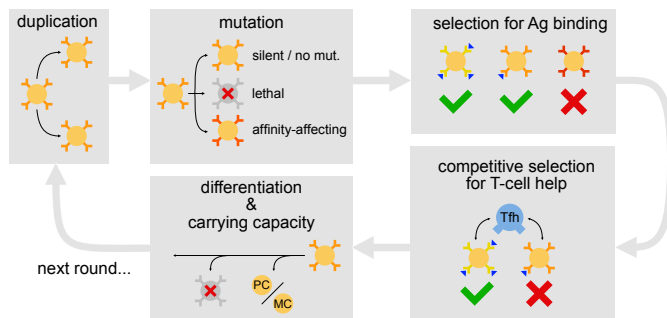


FIG. 1. Schematic representation of the processes that constitute an evolution round. Initially all cells in the population duplicate. Each daughter cell can develop affinity-affecting ($p_{aa} = 0.1$) or lethal ($p_{let} = 0.15$) mutations. Most of the affinity-affecting mutations have a deleterious effect. Cells are then selected on the base of their ability to bind the Antigen (Ag), and compete to receive T-cell help. Surviving cells have a small probability ($p_{diff} = 0.1$) of differentiating into Memory or Plasma Cells (MC/PC) and exiting the cycle. Moreover if the size of the population at this step exceeds the maximum carrying capacity ($N_{max} = 2500$) then extra cells are removed randomly. The remaining cells begin then a new evolution round.

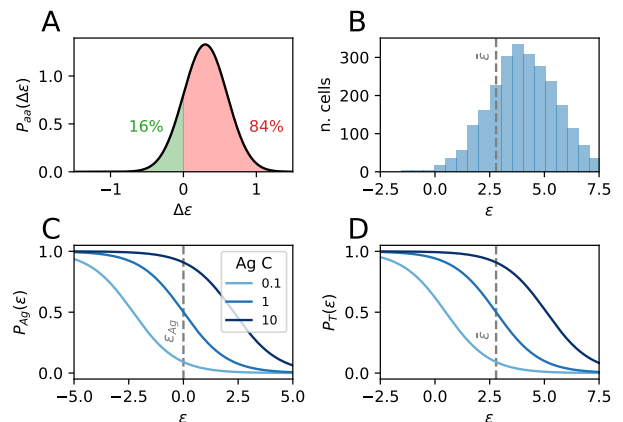


FIG. 2. **A**: probability distribution of the energy change $\Delta\epsilon$ introduced by affinity-affecting mutations. Most of the mutations are deleterious. **B**: example of binding energy histogram for the initial population of B-cells and the corresponding value of $\bar{\epsilon}$. **C**: probability for a cell with binding energy ϵ of surviving Ag-binding selection for different values of Ag concentration C . The threshold value ϵ_{Ag} is represented with a gray dashed line. An higher concentration corresponds to an higher survival probability. **D**: same as panel C, but for T-cell selection survival probability. Competition is introduced by the fact that in this case the threshold $\bar{\epsilon}$ depends on the binding energy distribution of the population, as shown in panel B.

by some amount $\Delta\epsilon$. We assume that $\Delta\epsilon$ is a random variable, extracted from a Gaussian distribution with mean μ_M and standard deviation σ_M (see fig. 2 A);

- not mutate or develop silent mutations, with probability p_{sil} . In both cases its affinity is unchanged;
- be hit by a lethal mutation with probability p_{let} , in which case it is removed from the population.

As a result, the distribution of the changes $\Delta\epsilon$ is therefore given by the kernel

$$K(\Delta\epsilon) = \frac{p_{aa}}{\sqrt{2\pi\sigma_M^2}} \exp\left(-\frac{(\Delta\epsilon - \mu_M)^2}{2\sigma_M^2}\right) + p_{sil} \delta(\Delta\epsilon) \quad (1)$$

where $\delta(\Delta\epsilon)$ is Dirac delta distribution. Notice that, due to lethal mutations, the integral of the kernel K is not normalized to unity but to $p_{aa} + p_{sil} = 1 - p_{let}$. Parameters are chosen such that only a small fraction of the mutations is beneficial, i.e. decreases the binding energy (cf. appendix A).

After duplication and mutation cells migrate to the LZ where they try to bind the Ag exposed on the surface of FDCs. Failure to do so results in cell death, and only cells that are able to bind the Ag with sufficient affinity survive this step of selection. Similarly to [13, 19] we consider the survival probability for a cell with binding

energy ϵ to be given by the following Langmuir isotherm:

$$P_{\text{Ag}}(\epsilon) = \frac{C e^{-\epsilon}}{C e^{-\epsilon} + e^{-\epsilon_{\text{Ag}}}}, \quad (2)$$

where ϵ_{Ag} is a threshold binding energy and C represents the dimensionless concentration of Ag available for cells to bind. This concentration controls the strength of selection, making successful binding more likely when more Ag is available to bind. In practice it acts by imparting a shift of magnitude $\log C$ to the energy threshold. The functional dependence of the selection probability on ϵ and C is displayed in fig. 2C.

In a second selection step cells compete to receive a survival signal from T-follicular helper cells, with the signal being preferentially delivered to cells that bind more Ag. The survival probability for a cell with binding energy ϵ is:

$$P_{\text{T}}(\epsilon, \bar{\epsilon}) = \frac{C e^{-\epsilon}}{C e^{-\epsilon} + e^{-\bar{\epsilon}}}, \quad \text{with } e^{-\bar{\epsilon}} = \langle e^{-\epsilon} \rangle_{\text{pop}} \quad (3)$$

Where the term $\langle e^{-\epsilon} \rangle_{\text{pop}}$ represents the average of this quantity over the population and encodes for competition, cf. fig. 2 B and D. The surviving cells then can differentiate into plasma or memort cells with total probability p_{diff} . We do not keep track of these differentiated cells in the simulation.

After this step if the population size exceeds the maximum carrying capacity N_{max} cells are randomly removed until this threshold is met. The surviving cells start then the next round of evolution. The values of the model parameters are reported in table I, and discussed in appendix A.

B. Population bottleneck and lineage survival

Similarly to other AM models [18, 19], for standard parameter values the population initially undergoes a bottleneck state. This is caused by the strong selection pressure initially imposed by Ag-binding selection, which later relaxes if the average population energy reaches values $\langle \epsilon \rangle_{\text{pop}} \sim \epsilon_{\text{Ag}}$. By controlling the selection pressure (cf. eqs. (2) and (3)) Ag concentration also impacts the population survival probability.

As an illustration we report in fig. 3 the average evolution of 1000 stochastic simulations for three different values of the concentration C . For all three values the population size initially decreases under the combined effect of the two selection steps (fig. 3A). This decrease lasts for few turns, and is accompanied by a quick increase in average affinity (fig. 3B). At this point surviving populations are composed of few high-affinity cells, on which the main acting selection force is competitive selection in eq. (3). If this selection pressure is not too strong then the population will later expand and mature. Through a mechanism analogous to the one studied in [13] Ag concentration then controls the maturation speed, as can be

seen by comparing the time dependence of the average energy in the cases $C = 10$ and $C = 2.7$ in fig. 3B.

The fraction of surviving simulations as a function of time is shown in fig. 3C. At low concentration ($C = 1$) the population goes quickly extinct in all simulations. For such small value of Ag concentration competitive selection alone is sufficiently strong to impede population growth. Intermediate concentration value ($C = 3.5$), on the contrary, are sufficient to sustain population growth. In this case extinction can nevertheless occur close to the bottleneck state, when population size gets transiently small, see fig. 3A; if some cells are able to survive and pass this bottleneck, then the population again grows to full size and continues maturation. Last of all, at high concentration ($C = 10$), the bottleneck pressure is not sufficient to significantly endanger population survival, and all simulations are able to overcome the low-population state without going extinct, but maturation proceeds very slowly. We will study in detail in the next section the dependence of the survival probability of the population of cells upon Ag concentration.

Survival and future expansion is also strongly dependent on the initial distribution of affinities. This effect can be readily observed on lineages originated from a single ancestor, with energy ϵ . In fig. 4 we display three examples of lineage evolution in the form of trees in which each node corresponds to a different cell. These lineages differ by the affinity of their progenitor at the root of the tree. The progeny of the lowest-affinity one (red, $\epsilon_i \sim -0.3$) goes extinct in few evolution rounds. In the one with intermediate affinity (orange, $\epsilon_i \sim -0.45$) only few individuals are able to survive the bottleneck. The high-affinity one (green, $\epsilon_i \sim -1.3$) instead expands and eventually takes over the population. To quantify the population survival probability we will first investigate how the survival of single lineages depends on the progenitor affinity.

III. PROBABILITY OF SURVIVAL AND DISTRIBUTION OF EXTINCTION TIMES

In this Section we study the probability that a B-cell lineage descending from a single progenitor cell survives through a population bottleneck, in particular how the probability of survival depends on the affinity of the progenitor. We also determine the distribution of extinction times of the lineage. We then make use of these results to evaluate the survival probability for the full population.

A. Case of one lineage

1. Probability of survival

Let us consider a progenitor cell with binding energy ϵ , present in the population at the beginning of evolution $t = 0$. At each evolution round this cell will divide and

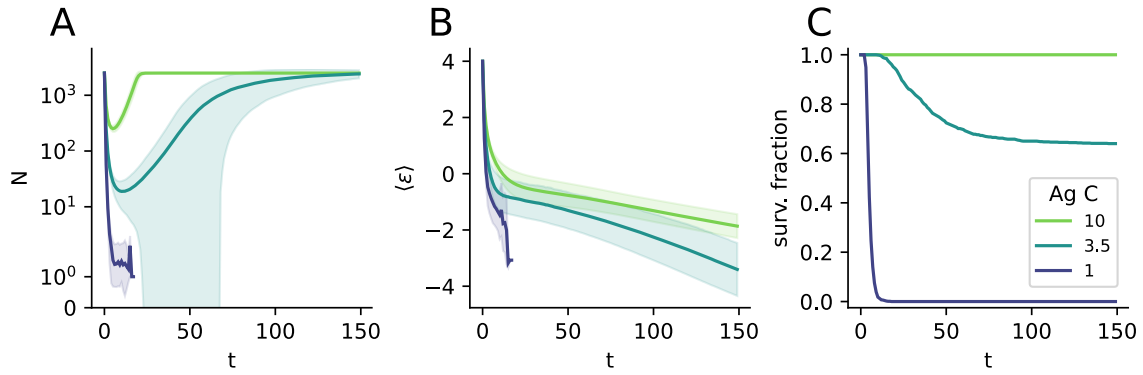


FIG. 3. Average evolution for 1000 different stochastic simulations of the model at three different levels of Ag concentration $C = 1, 3.5, 10$, color-coded according to the legend on the right. **A**: population size N as a function of evolution round. Shaded area covers one standard deviation for surviving simulations. The minimum population size on the bottleneck depends strongly on Ag concentration **B**: same as panel A but for the average population binding energy $\langle \epsilon \rangle$. Notice how for surviving populations the maturation speed depends on Ag concentration. **C**: Fraction of surviving simulations as a function of time. At low concentration the bottleneck drives all simulations to extinction, while at high concentration the population survives with high probability.

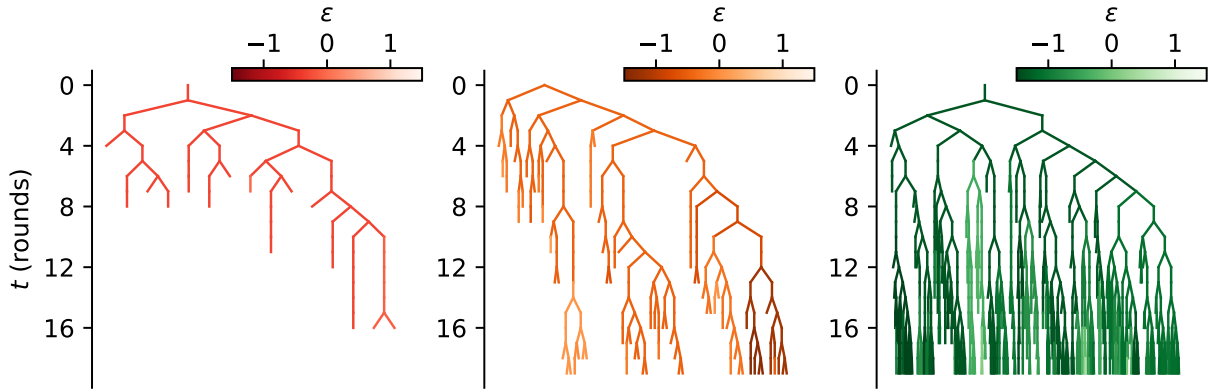


FIG. 4. Examples of stochastic lineage evolution through a population bottleneck. We perform a single simulation of our model at Ag concentration $C = 5$ and consider three different progenitors with different initial affinities (red $\epsilon_i \sim -0.3$, orange $\epsilon_i \sim -0.45$ and green $\epsilon_i \sim -1.3$). We represent their progeny evolution in the form of a tree with each cell corresponding to a node, and encoding affinity in the branch color. The lineage of the red progenitor quickly goes extinct, while the lineage of the orange progenitor survives the bottleneck but only with few individuals. The green progenitor lineage conversely survives the population bottleneck and undergoes great expansion. Notice how fate correlates with the initial progenitor affinity.

its offspring will have some probability of being removed from the population, either due to selection, differentiation or lethal mutations. In fig. 5A we report an example of lineage evolution for a progenitor with binding energy $\epsilon = 1$. Color indicates the binding energy of each cell, according to the color-scale on top. In this example cells accumulate deleterious mutations until the lineage eventually goes extinct after $t = 26$ evolution rounds.

We are interested in computing the probability $d_t(\epsilon)$ that *all of the offspring* of a progenitor with binding energy ϵ will be extinct by evolution round t , see fig. 5B. The expression for $t = 1$ can easily be written as the probability that both daughter cells generated during the duplication phase will be removed by the end of the round. As stated above, this can occur either by lethal muta-

tion, by failing selection or by differentiation. For each daughter cell this probability is more easily expressed as one minus the probability of not being removed:

$$d_1(\epsilon) = \left[1 - \int d\Delta \epsilon K(\Delta \epsilon) P_S(\epsilon + \Delta \epsilon) (1 - p_{\text{diff}}) \right]^2, \quad (4)$$

where the expression for $K(\Delta)$ is the one given in eq. (1), and $P_S(\epsilon)$ is the probability for a cell with binding energy ϵ of surviving selection. In the bottleneck state most of the selection pressure is generated by Ag-binding selection (i.e. $\bar{\epsilon}_t < \epsilon_{\text{Ag}}$). As a first approximation we therefore neglect competitive selection for T-cell help, and consider simply $P_S(\epsilon) = P_{\text{Ag}}(\epsilon)$ (cf. eq. (2)). This introduces two important simplifications. First, the expression of P_S does not depend on time. Second, removing the com-

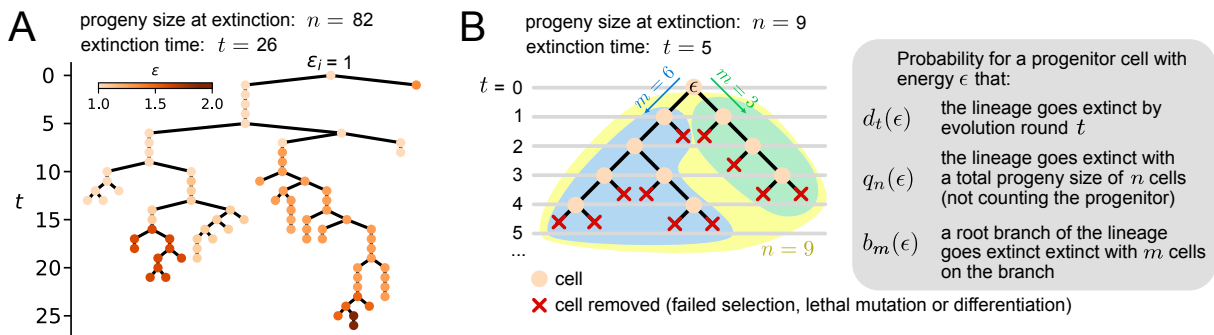


FIG. 5. **A.** Example of lineage issued from a progenitor with binding energy $\epsilon_i = 1$ obtained from a stochastic simulation performed at Ag concentration $C = 7$ in the approximation of only Ag-binding selection. Each node in the tree represents a cell, its binding energy ϵ encoded using the colorscale on top. In this example cells progressively accumulate deleterious mutations until after 26 evolution rounds the population eventually goes extinct. **B.** Schematic illustration of the quantities analyzed in our theory. On the left we depict a lineage evolution, stemming from a progenitor with binding energy ϵ . The probability that such a lineage goes extinct by time t is indicated with $d_t(\epsilon)$ (in this example $t = 5$). The quantity $q_n(\epsilon)$ represents instead the probability that the lineage goes extinct counting a total of n cells (not counting the progenitor) ($n = 9$ in this example). Finally, $b_m(\epsilon)$ is the probability that one of the two sub-lineage stemming from the progenitor goes extinct counting m cells (here $m = 6$ for the left branch and $m = 3$ for the right one).

petitive selection decouples the fate of all cells in the population.

The probabilities $d_t(\epsilon)$ for $t > 1$ can be evaluated using recursive relations that express the probability of extinction in t rounds as the probability for each daughter cell to either go extinct in one round, or to survive the first round but to have their respective offspring go extinct in $t - 1$ rounds:

$$d_t(\epsilon) = \left[1 - \int d\Delta \epsilon K(\Delta \epsilon) P_S(\epsilon + \Delta \epsilon) (1 - p_{\text{diff}}) \times (1 - d_{t-1}(\epsilon + \Delta \epsilon)) \right]^2 \quad (5)$$

In other words, the probability that all of the offspring goes extinct in t rounds is the probability that each of the two daughter cells generated during the duplication phase of the first round independently are removed before the end of round t . Since division is symmetric this probability must be the same for each daughter cell, and is the probability inside the square brackets. Then, this probability that the daughter cell is removed is more easily expressed as one minus the probability that it survives, and some of its offspring also survives for other $t - 1$ rounds (term in the integral). In turn this is expressed as the probability of surviving mutation (with a potential energy change of entity $\Delta \epsilon$), selection and differentiation, multiplied by the probability that the daughter cell offspring does not go extinct in $t - 1$ rounds (term on the second line).

In fig. 6A we plot the behavior of $d_t(\epsilon)$ as a function of evolution round t and binding energy ϵ (orange curves, color indicates extinction round t). As expected, the extinction probability is a monotonically increasing function of time and of energy, and reaches an asymptotic value $d_\infty(\epsilon)$ for large t . Our analytical result is in ex-

cellent agreement with simulations for the mean extinction probability (blue dots). The asymptotic probability $d_\infty(\epsilon)$ ranges between

- $d_\infty(\epsilon \rightarrow +\infty) = 1$, as high-energy, i.e. low-affinity cells do not pass the selection step and quickly go extinct, and
- $d_\infty(\epsilon \rightarrow -\infty) = \min\{1, [1 - 1/\alpha]^2\}$, with $\alpha = (1 - p_{\text{let}})(1 - p_{\text{diff}})$. The value of d_∞ for very high-affinity cells may be lower than one since lethal mutations and differentiation may drive the lineage to extinction, especially during the first few evolution rounds when the number of offspring is still small. The above expression for $d_\infty(\epsilon \rightarrow -\infty)$ can be obtained by searching a fixed point to eq. (5), and considering that, for $\epsilon \rightarrow -\infty$, mutations do not sensibly change the survival probability and can be neglected. The parameter α defined above is then the probability for a high-affinity cell to survive one round and not be removed by lethal mutations or through differentiation. Notice that, if $\alpha < 1/2$, we have $d_\infty(\epsilon \rightarrow +\infty) = 1$, as is to be expected when on average less than one individual in the offspring will survive. However this case is pathological: in this regime, irrespective of the progenitor energy, the population always goes extinct ($d_\infty(\epsilon) = 1$).

2. Distribution of extinction times

The probability that a lineage generated by a progenitor with energy ϵ goes extinct exactly at round t can easily be expressed as

$$r_t(\epsilon) = d_t(\epsilon) - d_{t-1}(\epsilon). \quad (6)$$

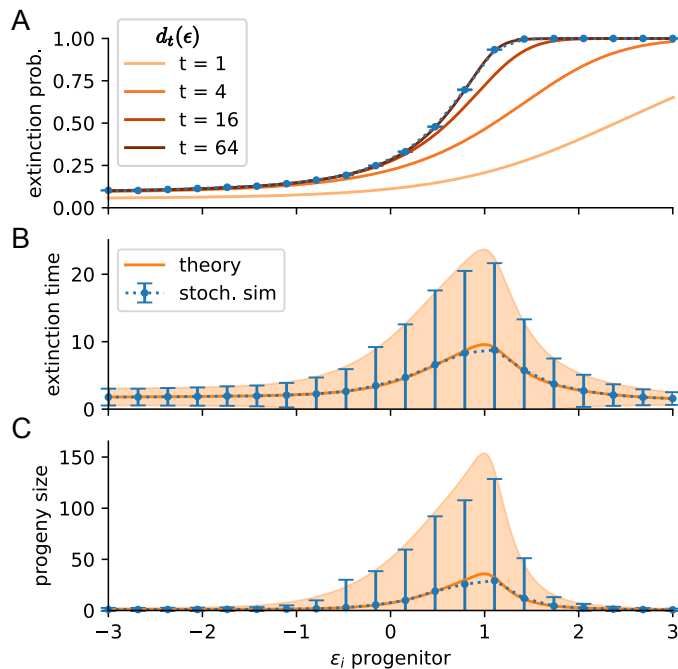


FIG. 6. Comparison between stochastic simulations (blue) and theory (orange) for the probability of extinction (A), lineage extinction time (B) and average progeny size at extinction (C) as a function of the progenitor energy ϵ_i in absence of competitive selection. For each conditions we consider 5000 different stochastic simulations that terminate with extinction at Ag concentration $C = 7$. **A:** stochastic extinction probability (blue dots, error bar indicate the standard error of the mean) evaluated as the fraction of simulations that terminate with extinction over the total number of simulations performed. This is compared to the value of $d_t(\epsilon)$ as described by our theory. **B:** mean and standard deviation of extinction time (blue) over 5000 simulations terminating in extinction. This is compared to the theoretical prediction (orange) for the mean and standard deviation of this quantity, obtained using the time extinction probability $r_t(\epsilon)$. **C:** same as B but for the progeny size. In this case the theoretical predictions are obtained using the generating function theory.

This allows us to evaluate the mean and variance for the extinction time probabilities (see fig. 6B) simply from the first two moments of the distribution:

$$\langle t \rangle_\epsilon = \sum_{t=0}^{\infty} t r_t(\epsilon), \quad \langle t^2 \rangle_\epsilon = \sum_{t=0}^{\infty} t^2 r_t(\epsilon), \quad (7)$$

In fig. 6B we compare, in the approximation of no competitive selection, the average extinction time computed from simulations (blue, error bars indicate the standard deviation of extinction times for each progenitor affinity) with theoretical predictions (orange, shaded area covers one standard deviations). We again find a very good match. The average extinction time shows a peak for intermediate affinities, which can be interpreted as follows. Low-affinity progenitors, i.e. having high binding energy have close-to-one probability of extinction, and very often go extinct in the first few rounds. High-affinity cells on the contrary have a small but non-zero probability of extinction, see value of extinction probability in fig. 6A. This is mainly due to affinity-independent terms such as differentiation probability, which confer to the lineage a small chance of going extinct after the first selection rounds, when the progeny is still small. For affinities close to $\epsilon = 1$ we observe intermediate values for the probability of survival, and maximum value for

average extinction time. This behaviour can be better understood when mutations are turned off, in which case equations can be solved exactly, as shown below.

3. Exactly solvable case of no mutation

We hereafter consider the case of no affinity-affecting mutations, for which the mutation kernel eq. (1) reads

$$K(\Delta) = (1 - p_{\text{let}}) \delta(\Delta). \quad (8)$$

In this case genealogies belong to the class of Galton-Watson trees [21], and the asymptotic survival probability can be derived exactly. This probability is better expressed by considering the quantity

$$\gamma(\epsilon) = (1 - p_{\text{let}}) P_S(\epsilon) (1 - p_{\text{diff}}) \quad (9)$$

which represents the probability for a daughter cell to remain in the GC and not be removed by either lethal mutations, selection or differentiation. The infinite-time extinction probability $d_\infty(\epsilon)$ can be found by rewriting eq. (5) in the limit $t \rightarrow \infty$:

$$d_\infty(\epsilon) = \min \left(1, \left(\frac{1}{\gamma(\epsilon)} - 1 \right)^2 \right). \quad (10)$$

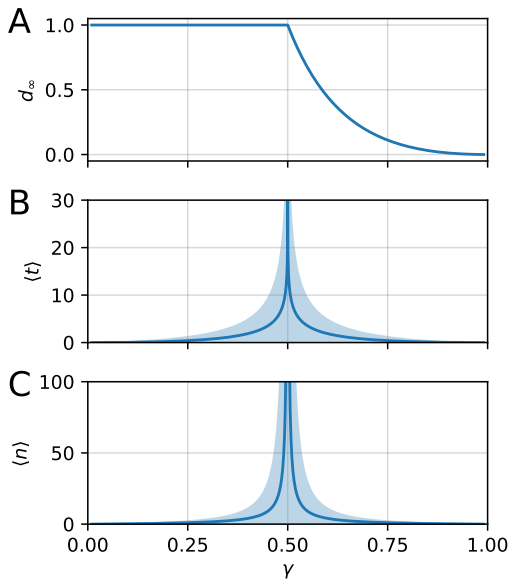


FIG. 7. Value of the extinction probability d_∞ (A), average extinction time $\langle t \rangle$ (B) and average progeny size $\langle n \rangle$ (C) as a function of the survival probability $\gamma(\epsilon)$ in (9) in the approximation of no affinity-affecting mutation. In B and C shaded area covers one standard deviation. Notice how extinction times and genealogy sizes diverge at $\gamma = \frac{1}{2}$.

As expected, lineages will always go extinct if the average number of surviving offspring at division is not greater than one: $d_\infty(\epsilon) = 1$ if $\gamma(\epsilon) \leq \frac{1}{2}$. In fig. 7A we report the behavior of d_∞ as a function of γ . This function presents a singularity at the critical value $\gamma = 1/2$, for which the Galton-Watson process is critical.

Finding an explicit expression for the distribution of extinction times is harder, but results can be obtained for the critical value $\gamma = \frac{1}{2}$. We find that the extinction time probability behaves asymptotically as a power law, with infinite mean and variance: $r_t \sim 4/t^2$ for large t . This result, which is a known feature of critical Galton-Watson processes, can be verified by inserting the Ansatz $d_t \sim 1 - \alpha t^{-1} + o(t^2)$ in eq. (5), together with the simplified form of the mutation kernel (8) and the assumption that $\gamma = 1/2$. The only admissible solution is $\alpha = 4$ which, combined with the definition of r_t eq. (6), proves the statement. In fig. 7B we display the mean and variance of the extinction time distribution as a function of γ . Comparison with fig. 6B shows that the divergence is removed when evolution includes affinity-affecting mutations. Mutations drive lineages away from the critical line, either to high affinities and survival, or to lower affinities and extinction.

B. Case of full population

Building on the results derived above, we now turn to the problem of quantifying the average probability of

extinction for the whole population.

As a first approximation we do not consider competitive selection, since most of the selection pressure in a bottleneck is given by Ag-binding selection. Given a total of N_i cells in the initial population, having energies $\{\epsilon_k\}_{k=1\dots N_i}$ the probability that all cells will be extinct by evolution round t is simply given by the product of extinction probabilities for all cells $\prod_k d_t(\epsilon_k)$. Moreover, since the initial energies are independently extracted from a Gaussian distribution $\varphi(\epsilon)$ with mean μ_i and standard deviation σ_i , the average extinction probability by round t over all possible extractions of the initial population is given by:

$$P_{\text{ext}}(t) = \left(\int d\epsilon \varphi(\epsilon) d_t(\epsilon) \right)^{N_i} \quad (11)$$

With the help of this formula we evaluate the average survival probability as a function of Ag concentration C and initial population size N_i , and compare the prediction with stochastic simulations in which we turn off T-cell selection. The results, reported in fig. 8B and C (blue), match exactly.

In the presence of competitive selection the empirical survival probability evaluated from simulations slightly decreases, compare blue and orange dotted lines in fig. 8B and C. The theory can be extended to account for T-selection in an effective manner. In practice, one needs first to extend the theory to include a time-dependence of the survival probability. At this point competitive selection can be included by introducing an effective coupling between cells in a ‘mean field’ fashion, by estimating the average evolution of the term $\bar{\epsilon} = -\log\langle e^{-\epsilon} \rangle_{\text{pop}}$ contained in the expression for the T-selection survival probability eq. (3).

Assume that the survival probability $P_S(\epsilon, t)$ is now time-dependent. The probability of extinction does not depend solely on the number of evolution rounds anymore, but also on the initial time at which the progenitor is considered. We define $d_{t,s}(\epsilon)$ as the probability that a cell, which at the end of round t has binding energy ϵ , will have all of its offspring extinct by the end of round $s > t$. For any value $t \geq 0$ we can write as before the probability of extinction in one round:

$$d_{t,t+1}(\epsilon) = \left(1 - \int d\Delta K(\Delta) P_S(\epsilon + \Delta, t) (1 - p_{\text{diff}}) \right)^2 \quad (12)$$

And for any pair of rounds $s > t \geq 0$, with $s - t > 1$, the following recursive relation, analogous to eq. (5), holds:

$$d_{t,s}(\epsilon) = \left[1 - \int d\Delta K(\Delta) P_S(\epsilon + \Delta, t) (1 - p_{\text{diff}}) \times (1 - d_{t+1,s}(\epsilon + \Delta)) \right]^2 \quad (13)$$

Finally, similarly to eq. (11), the probability that the full

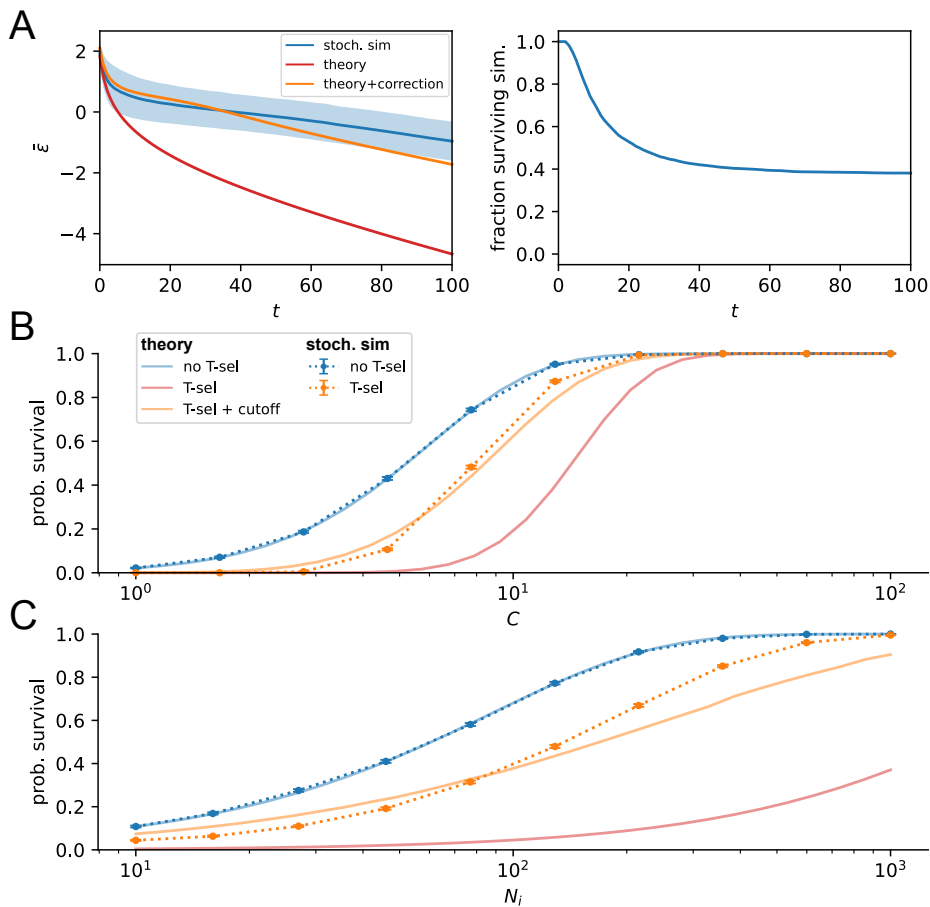


FIG. 8. Probability of population survival in a bottleneck condition as a function of initial population size N_i and Ag concentration C . **A**: Left: comparison between the evolution of $\bar{\epsilon}$ in stochastic simulations (blue, mean and standard deviation over 5000 simulations) and theoretical prediction without (red) and with (orange) finite-size correction. This correction consists in cutting the tail of the initial energy distribution in proximity of the expected value for the highest-affinity individual. The correction improves the prediction for the evolution of $\bar{\epsilon}$ at short times. Right: fraction of surviving simulations as a function of evolution round. With the finite-size correction the value of $\bar{\epsilon}$ is well-approximated during the time it takes for most of the simulations to go extinct. In this example we set $C = 7$, $N_i = 100$. **B**: bottleneck survival probability as a function of antigen concentration C . Comparison between stochastic simulations (dotted line, error bars indicate the standard error of the mean) and the predictions our theory (full lines). Stochastic simulations are reported both without (blue) and with (orange) competitive selection for T-cell help (T-sel). For the theory instead we consider the case without T-sel (blue), with T-sel (red) and with T-sel plus finite-size correction (orange). In the absence of T-sel all cells evolve independently, and the theory and simulations match exactly. The inclusion of T-sel slightly decreases the survival probability in stochastic simulations. Accounting for this contribution by using the infinite-size estimate for the evolution of $\bar{\epsilon}$ overestimates the selection pressure. Adding the finite-size correction results in a much better estimate. In this example we set $N_i = 100$. **C**: same as B, but the survival probability is evaluated as a function of the initial size of the population N_i . Here we set $C = 7$.

population goes extinct by evolution round t is given by:

$$P_{\text{ext}}(t) = \left(\int d\epsilon \varphi(\epsilon) d_{0,t}(\epsilon) \right)^{N_i}, \quad (14)$$

where $\varphi(\epsilon)$ is a Gaussian distribution with mean μ_i and standard deviation σ_i .

At this point we can make explicit the time dependence of the survival probability including selection for T-cell help: $P_S(\epsilon, t) = P_{\text{Ag}}(\epsilon) P_T(\epsilon, \bar{\epsilon}_t)$ (cf. eq. (3)). Given the stochastic nature of our model, the variable $\bar{\epsilon}_t$ which quantifies selection pressure is in reality a stochas-

tic variable. We estimate its average evolution using the large-population-size limit described in appendix B, under which the model becomes deterministic. This allows us to numerically evaluate the extinction probability eq. (14). The outcome, however, underestimates the real survival probability (compare red curve and orange dotted line in fig. 8B and C). This mismatch originates mainly from the fact that in the big-size approximation $\bar{\epsilon}$ evolves faster than in stochastic simulations (cf. blue and orange line in fig. 8A-left). In turn, this occurs because the value of $\bar{\epsilon}$ is strongly dependent on the high-affinity

tail of the population, whose evolution is influenced by finite-size effects.

This discrepancy can, however, be reduced with a simple finite-size correction. This correction is based on the consideration that the large-size limit of the model approximates the population binding energy histogram with a continuous distribution, encoded in the density function $\rho_t(\epsilon)$ (cf. appendix B). At the beginning of evolution this function takes the shape of a normal distribution, corresponding to the initial binding energy distribution of naive responders, with tails extending indefinitely in both directions. As the population is finite in reality, consisting of N_i individuals, we do not expect these tails to be populated. The correction procedure consists in removing these tails, by setting the initial distribution equal to zero outside a range delimited by two values $[\epsilon^-, \epsilon^+]$.

These two values are chosen equal to the expected energy of, respectively, the highest and lowest affinity individual in the population. The probability distribution for their binding energies can be expressed as a function of the naive binding energy distribution $\varphi(\epsilon)$ (as before a Gaussian with mean μ_i and variance σ_i^2) from which the energy of all cells is extracted. If we call $F(\epsilon) = \int_{-\infty}^{\epsilon} d\epsilon' \varphi(\epsilon')$ the cumulative distribution function, then these distributions can be expressed as:

$$\varphi^+(\epsilon) = \frac{d}{d\epsilon} [F(\epsilon)]^{N_i} \quad (15)$$

$$\varphi^-(\epsilon) = -\frac{d}{d\epsilon} [1 - F(\epsilon)]^{N_i} \quad (16)$$

The values ϵ^\pm simply correspond to the means of these distributions.

Removing the tails to the initial distribution causes an initial slow-down in the evolution of $\bar{\epsilon}$ (cf. green line in fig. 8A-left). This slow-down is eventually lost, but the agreement remains for a time sufficient for most of the stochastic simulations to go extinct (cf. fig. 8A-right) which is the relevant timescale to capture bottleneck survival.

Taking the value of $\bar{\epsilon}$ obtained by combining the big-size approximation (cf. appendix B) with the cutoff correction described above, and using it to evaluate the population survival probability, we obtain a much better agreement of the theory with simulations (compare orange curve and orange dotted line in fig. 8B and C). The remaining discrepancy are due to the fact that the average evolution of $\bar{\epsilon}$ is still not exactly captured, and the ‘mean-field’ nature of our approximation, which neglects the feedback of the energies in the population onto $\bar{\epsilon}$.

IV. LINEAGE SIZE AT EXTINCTION

In this Section we focus on the distribution of sizes of the progeny at extinction. This size strongly depends on

the model parameters, such as the energy of the progenitor. An example is displayed in fig. 5A, in which the lineage consists of a total of 82 cells. Like extinction time, this quantity is well-defined only for lineages that go extinct. Populations that are able to pass the bottleneck undergo exponential growth, with a rate that can be calculated from the large-size theory of Appendix B, see [13].

1. Recursion equations for the distribution of sizes

Similarly to what was done in the previous Section for the extinction time and probability, we now derive a recursive formula to quantify the total offspring size. We need to keep track of the sum of two random variables representing the numbers of descendants of each daughter cell. The recursion therefore includes a convolution, which is numerically harder to compute but can be handled using probability generating functions. The recursive relations can be expressed in term of these functions, and can be used to evaluate the moments of the probability distribution without having to numerically perform the convolution.

We name $q_n(\epsilon)$ the probability that a progenitor with energy ϵ generates a total offspring of exactly n cells before extinction, not counting the progenitor itself (see fig. 5B). This probability can be better expressed if we separate the contribution of the two daughter cells. Considering genealogies encoded as binary trees, we call $b_m(\epsilon)$ the probability that along the branch corresponding to one of the daughter cells of a progenitor with energy ϵ we find a total of m descendants (including the daughter cell itself) before extinction (see fig. 5). The expression for $m = 0$ is simply given by the probability that the daughter cell is removed before the end of the round:

$$\begin{aligned} b_0(\epsilon) &= 1 - \int d\Delta K(\Delta) P_S(\epsilon + \Delta) (1 - p_{\text{diff}}) \\ &= \sqrt{d_1(\epsilon)} \end{aligned} \quad (17)$$

The recursive relation in this case is composed of two equations. The first is a convolution that decomposes the probability of having n descendants as a sum over all possible repartitions of the descendant number along the two branches:

$$q_n(\epsilon) = \sum_{m=0}^n b_m(\epsilon) b_{n-m}(\epsilon) \quad (18)$$

The second expresses the probability to find m descendants along a branch as the probability that the daughter cell survives and has $m - 1$ descendants:

$$b_m(\epsilon) = \int d\Delta K(\Delta) P_S(\epsilon + \Delta) (1 - p_{\text{diff}}) q_{m-1}(\epsilon + \Delta) \quad (19)$$

We introduce the generating functions $Q(z, \epsilon)$ and $B(z, \epsilon)$, defined as:

$$Q(z, \epsilon) = \sum_{n=0}^{\infty} q_n(\epsilon) z^n, \quad B(z, \epsilon) = \sum_{m=0}^{\infty} b_m(\epsilon) z^m. \quad (20)$$

In terms of these generating functions equations eqs. (18) and (19) become

$$Q(z, \epsilon) = B(z, \epsilon)^2 \quad (21)$$

$$\frac{1}{z} [B(z, \epsilon) - b_0(\epsilon)] = \int d\Delta K(\Delta) P_S(\epsilon + \Delta) \times (1 - p_{\text{diff}}) Q(z, \epsilon + \Delta) \quad (22)$$

These relations can be used to evaluate the moments of these distributions with two additional considerations. The first is that $\sum_{n=0}^{\infty} q_n(\epsilon) = d_{\infty}(\epsilon)$. This sum does not converge to one since it only considers lineages that eventually go extinct. For the functions Q and B this translates into:

$$Q(z = 1, \epsilon) = d_{\infty}(\epsilon), \quad B(z = 1, \epsilon) = \sqrt{d_{\infty}(\epsilon)}. \quad (23)$$

Secondly, the moments of the distributions can be evaluated from the generating functions as:

$$\langle n^k \rangle_{\epsilon} = \frac{1}{d_{\infty}(\epsilon)} \sum_{n=0}^{\infty} n^k q_n(\epsilon) \quad (24)$$

$$= \frac{1}{d_{\infty}(\epsilon)} (z \partial_z)^k Q(z, \epsilon)|_{z=1}$$

$$\langle m^k \rangle_{\epsilon} = \frac{1}{\sqrt{d_{\infty}(\epsilon)}} \sum_{m=0}^{\infty} m^k b_m(\epsilon) \quad (25)$$

$$= \frac{1}{\sqrt{d_{\infty}(\epsilon)}} (z \partial_z)^k B(z, \epsilon)|_{z=1}$$

Applying the operator $z \partial_z$ one and two times on eq. (21) restitutes the following relations between the first two moments:

$$\langle n \rangle_{\epsilon} = 2 \langle m \rangle_{\epsilon}, \quad \langle n^2 \rangle_{\epsilon} = 2 \langle m^2 \rangle_{\epsilon} + 2 \langle m \rangle_{\epsilon}^2 \quad (26)$$

This corresponds simply to the fact that the total number of descendants is the sum of the descendants along the two branches. Applying the same operator on eq. (22) gives:

$$\sqrt{d_{\infty}(\epsilon)} \langle m \rangle_{\epsilon} = \int d\Delta K(\Delta) P_S(\epsilon + \Delta) (1 - p_{\text{diff}}) \times d_{\infty}(\epsilon + \Delta) [2 \langle m \rangle_{\epsilon + \Delta} + 1] \quad (27)$$

$$\sqrt{d_{\infty}(\epsilon)} \langle (m-1)^2 \rangle_{\epsilon} = \int d\Delta K(\Delta) P_S(\epsilon + \Delta) \times (1 - p_{\text{diff}}) d_{\infty}(\epsilon + \Delta) [2 \langle m^2 \rangle_{\epsilon + \Delta} + 2 \langle m \rangle_{\epsilon + \Delta}^2 + 1] \quad (28)$$

These equations can be solved numerically if we express them as fixed-point equations for the functions $\langle m \rangle_{\epsilon}$ and

$\langle m^2 \rangle_{\epsilon}$:

$$\langle m \rangle_{\epsilon} = \frac{1}{\sqrt{d_{\infty}(\epsilon)}} \int d\Delta K(\Delta) P_S(\epsilon + \Delta) (1 - p_{\text{diff}}) \times d_{\infty}(\epsilon + \Delta) [2 \langle m \rangle_{\epsilon + \Delta} + 1] \quad (29)$$

$$\langle m^2 \rangle_{\epsilon} = \frac{1}{\sqrt{d_{\infty}(\epsilon)}} \int d\Delta K(\Delta) P_S(\epsilon + \Delta) (1 - p_{\text{diff}}) \times d_{\infty}(\epsilon + \Delta) [2 \langle m^2 \rangle_{\epsilon + \Delta} + 2 \langle m \rangle_{\epsilon + \Delta}^2 + 4 \langle m \rangle_{\epsilon + \Delta} + 1] \quad (30)$$

The moments for n can then easily be evaluated using eq. (26).

In fig. 6C we compare the theoretical prediction for the first two moments (orange line represents the mean and shaded area covers one standard deviation) with the corresponding quantities from stochastic simulations (blue, error bars cover one standard deviation). Once more we find a good match. The peak at intermediate affinities can be explained, as done above for the extinction time, considering the critical nature of this phenomenon at intermediate values of the binding energy. This is done in the next section.

2. Exactly solvable case of no mutation

Similarly to what done for extinction probability, in the absence of affinity-affecting mutations we can find an explicit expression for the mean and variance of the population size at extinction. By plugging the simplified expression for the mutation kernel eq. (8) into eqs. (21) and (22) one obtains the following second degree equation for the generating function B :

$$z \gamma(\epsilon) B(z, \epsilon)^2 - B(z, \epsilon) + 1 - \gamma(\epsilon) = 0 \quad (31)$$

Where as before $\gamma(\epsilon)$ is the probability for a daughter cell not to be removed from the population during the evolution round, cf. eq. (9). This equation has two solutions. The correct one can be chosen by considering that B must be a monotonically increasing function of z . This gives:

$$B(z, \epsilon) = \frac{1 - \sqrt{1 - 4z \gamma(\epsilon) (1 - \gamma(\epsilon))}}{2z \gamma(\epsilon)} \quad (32)$$

The function $Q(z, \epsilon)$ can be evaluated from eq. (21), and the mean and variance for the population extinction sizes can be obtained using eq. (24). This results in:

$$\langle n \rangle_{\epsilon} = \begin{cases} \frac{2\gamma(\epsilon)}{1-2\gamma(\epsilon)} & \text{if } \gamma(\epsilon) < 1/2 \\ \frac{2-2\gamma(\epsilon)}{2\gamma(\epsilon)-1} & \text{if } \gamma(\epsilon) > 1/2 \end{cases} \quad (33)$$

$$\langle n^2 \rangle_{\epsilon} - \langle n \rangle_{\epsilon}^2 = \langle n \rangle_{\epsilon} (\langle n \rangle_{\epsilon} + 1) (\langle n \rangle_{\epsilon} + 2) / 2 \quad (34)$$

This quantity are reported as a function of γ in fig. 7C. Similarly to what observed for the mean and variance of the extinction time, Both of these quantities diverge for

the critical value $\gamma = \frac{1}{2}$, but this divergence is removed when mutations are considered.

It is interesting to consider the effect of this divergence on the coefficients $q_n(\epsilon)$, that represent the probability of a lineage that stems from a progenitor with binding energy ϵ to go extinct with a total progeny of n cells (see fig. 5B). From the definition of the generating function $Q(z, \epsilon)$ (cf. eq. (20)) it follows that the coefficients q_n can be obtained by Taylor expansion of this function around $z = 0$. In turn Q can be easily obtained from eq. (32) using the property $Q = B^2$ (see eq. (21)). The expansion results in the following expression for the coefficients when mutations are absent:

$$q_n = \frac{(2n+2)!}{(n+1)!(n+2)!} \gamma^n (1-\gamma)^{n+2} \stackrel{n \gg 1}{\approx} \frac{(1-\gamma)^2}{\sqrt{\pi n^3}} (4\gamma(1-\gamma))^n \quad (35)$$

In general these probabilities decay exponentially fast as a function of the size n . At the critical value $\gamma = 1/2$ however the term $4\gamma(1-\gamma)$ becomes equal to 1 and the coefficients decay as power law, with $n^{-3/2}$. This same result can more simply be obtained by considering that for any fixed value of ϵ the asymptotic decay of the coefficients q_n for $n \gg 1$ is related to the behavior of the generating function Q around its singularity at $z_c = 1/(4\gamma(1-\gamma))$. In particular in our case $Q \propto (1 - z/z_c)^{1/2}$ close to the singularity, and therefore $q_n \propto n^{-3/2} z_c^{-m}$ [22], which gives the expected asymptotic behavior described above.

3. Case of small-effect mutations

Corrections to the asymptotic behavior above arise when small-effect mutations are considered. In particular in eq. (8) we substitute the Dirac delta distribution with a peaked Gaussian, and consider a mutation kernel having the form:

$$K(\Delta) = (1 - p_{\text{let}}) \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{\Delta^2}{2\sigma^2}\right) \quad \text{with } \sigma \ll 1 \quad (36)$$

When the standard deviation is small enough one can approximate the integrals in eqs. (17) and (22) by Taylor-expanding the functions that multiply the mutation kernel around $\Delta = 0$. This results in the following approximation for eq. (22):

$$B - 1 + \gamma + \frac{\sigma^2}{2} \gamma'' = z\gamma B^2 + z \frac{\sigma^2}{2} [\gamma B^2]'' \quad (37)$$

according to the definition of γ (cf. eq. (9)), and with inverted commas indicating derivatives with respect to ϵ . This equation is analogous to eq. (31) with the addition of perturbation terms of the order of σ^2 . We therefore suppose $B(z, \epsilon) \sim B_0(z, \epsilon) + \sigma^2 \Delta B(z, \epsilon)$ where B_0 is given by eq. (32) and ΔB represents a perturbation to this $\sigma^2 = 0$ solution. By plugging this Ansatz in the previous equation we find the following expression for the

perturbation:

$$\Delta B = \frac{(B_0 - 1)}{2\gamma r^4} \left[2z\gamma'^2(1-2\gamma) + r \frac{\gamma'^2}{\gamma} + r^2 \left(\gamma'' - \frac{\gamma'^2}{\gamma} \right) \right], \quad (38)$$

where r is defined through

$$r = \sqrt{1 - 4z\gamma(1-\gamma)} = \sqrt{\frac{z_c - z}{z_c}}. \quad (39)$$

The critical value of z is given as before by $z_c = 1/(4\gamma(1-\gamma))$. Moreover, when the survival probability P_S is given by eq. (2), the derivatives of γ (cf. eq. (9)) can be expressed as

$$\gamma' = -\gamma(1-\tilde{\gamma}) \quad (40)$$

$$\gamma'' = \gamma(1-\tilde{\gamma})(1-2\tilde{\gamma}), \quad (41)$$

where $\tilde{\gamma} = P_{\text{Ag}}(\epsilon)$. From the expression for ΔB we can derive the perturbation to Q by using eq. (21). Keeping only the higher order in σ^2 , we obtain

$$Q \sim Q_0 + \sigma^2 \Delta Q, \quad \text{with } \Delta Q = 2 B_0 \Delta B \quad (42)$$

As stated in the previous section, the behavior of the probabilities q_n is strictly related to the behavior of Q around its singularity z_c . In particular the magnitude of the perturbation to the coefficients q_n introduced by mutations can be derived from the study of ΔQ . If we operate the substitution $z = z_c(1-r^2)$ and expand ΔQ in powers of r we obtain:

$$\Delta Q = c_4 r^{-4} + c_3 r^{-3} + c_2 r^{-2} + c_1 r^{-1} + O(1) \quad (43)$$

with the following expressions for the coefficients:

$$\begin{aligned} c_4 &= (1-2\gamma)^2(1-\tilde{\gamma})^2 \\ c_3 &= -(1-2\gamma)^2(1-\tilde{\gamma})^2 \\ c_2 &= -2(1-\gamma)(1-\tilde{\gamma})(1-2\gamma\tilde{\gamma}) \\ c_1 &= 2(1-\gamma)(1-\tilde{\gamma})(3-2\gamma-2\gamma\tilde{\gamma}) \end{aligned} \quad (44)$$

Based on this expansion we can derive an expression for the perturbation to the coefficients $\Delta q_n = q_n - q_n^0$ caused by weak mutations [22], where q_n^0 represents the value in the absence of mutations (cf. eq. (35)). We obtain that, for large n ,

$$\begin{aligned} \Delta q_n &= \sigma^2 [c_4(n+1) + c_3(n/\pi)^{1/2}(2+3/4n) + c_2 \\ &\quad + c_1(\pi n)^{-1/2} + O(n^{-3/2})] (4\gamma(1-\gamma))^n \end{aligned} \quad (45)$$

In fig. 9A we plot the values of the coefficients c as a function of $\tilde{\gamma}$. It is interesting to notice that both c_4 and c_3 , controlling the two leading orders, are null for the critical value $\gamma = 1/2$. This leaves the next leading order to c_2 which, for this value of γ , is negative. As a result, we expect that the perturbation tends to lower the values of q_n at large n for $\gamma \sim 1/2$, and to raise it for $\gamma \lesssim 1/2$. This means that large-size extinction events are made less probable by the presence of mutations around

the critical value, which is consistent with the removal of the divergence observed in fig. 7C when mutations are present.

In fig. 9B,C,D we compare the above prediction for Δq_n with the value provided by numerical simulations for different values of γ , as indicated in each plot. We expect the theoretical prediction to be accurate for large values of n , and as long as the perturbation Δq_n remains small with respect to the unperturbed value q_n^0 . However, for increasing values of n the perturbation grows faster than the unperturbed value, eventually invalidating this assumption. As a proxy for an accuracy upper limit we mark with a green dotted line the value of n at which the perturbation Δq has a magnitude equal to 10% of the unperturbed value q_n^0 . Based on eqn. (45) this value scales as $\sigma^{-4/5}$ for small σ 's.

V. DISCUSSION

In this work we focused on the effects of a bottleneck on a B-cell population in the course of the affinity maturation process. Through a recursive relation that links the probability of bottleneck survival of a cell to the one of its daughter cells we were able to retrieve the dependence of a lineage extinction probability on its progenitor affinity. For lineages that go extinct we also evaluated the mean and variance of extinction time and progeny size, revealing a peak in extinction time corresponding to average affinity progenitors. Lineages stemming from these progenitor spawn in equilibrium between extinction and survival, and persist in this state until mutations drive the lineage either to survival or to extinction. Building on these results we then evaluated the survival probability for the full population as a function of Ag concentration and population size. We also included the effect of competition in an effective manner, using the deterministic model limit combined with a finite-size correction.

The bottleneck phenomenology was included in different maturation models [18, 19], which considered as optimal the maturation regime in which the B-cell population was subject to a strong enough selection force to grant good affinity enhancement, while at the same time not strong enough to cause population extinction. While the properties of the above models were numerically evaluated based on stochastic simulations, we here present exact or approximate derivations for various quantities of interest through the use of recursive relations and probability generating functions. These techniques are often encountered in the context of branching processes, and similar approaches have been used in the study of AM [23, 24]. In our case we coupled the theoretical analysis with a more realistic model that included both the effect of mutations and selection for Ag binding and competition. We hope our approach both provides a better understanding of what controls the lineage survival probability, and requires less computational resources when compared to averaging over many stochastic simulations.

One could therefore easily explore the effect of changing different model parameters on the survival probability and the lineage size.

Compared to real AM, our model is obviously simplified in many aspects, for example we do not impose an affinity ceiling and beneficial mutations can accumulate indefinitely. However we believe this approximation to be reasonable since for the bottleneck scenario we consider low-affinity cells that could potentially undergo many affinity-improving mutations. Moreover we consider Ag concentration to be constant, while in reality Ag is subject to natural decay and consumption by B-cells. We believe this approximation to be acceptable when studying bottleneck survival, that in most cases is resolved in few evolution rounds. In our analysis we focused on the evolution of the population of B-cells inside a single GC. In the course of a response however many GCs form inside the body.² By averaging over the distribution of energies in the initial population our theory indeed quantifies the average bottleneck survival probability of GCs in this ensemble, in the hypothesis that the affinity of cells each initial population is independent. If instead affinities are correlated or if cells are able to migrate between GCs then the average would be harder to compute. Unfortunately the lack of precise experimental quantification of these processes forbids any meaningful modeling so far.

An important implication of our work is the relation between progenitor affinity to lineage survival probability. One of the current challenges in vaccine design against mutable pathogens consists in designing immunogens that are capable of stimulating specific naive B-cell precursors [26]. The affinity of these naive precursors for the immunogens, along with the precursors frequency, are important variables for the successful colonization of GCs by these lineages [27]. In this context our theoretical work defines a natural affinity threshold in proximity of the extinction time peak (cf. fig. 6B). Designed immunogens in vaccines should aim at binding a target progenitor cell with affinity around this peak, and not much lower. Theory suggests that the progeny of such a cell will reside in the GC for a long enough time to have a non-negligible chance of accumulating enough mutations to increase the affinity for the Ag. A progenitor with lower affinity instead might be removed too quickly from the GC to be able to accumulate sufficient mutations.

The results obtained here could be extended in different directions. For instance it would be interesting to investigate how stochasticity and permissiveness in selection influence survival probability. These effect, which are supposed to play a major role in AM against complex pathogens [13, 28], might confer an affinity-independent contribution to the survival probability and

² While their number is not known with accuracy, it could range from many tens to few hundreds since spleen sections revealed around 20-50 GCs in mice [25].

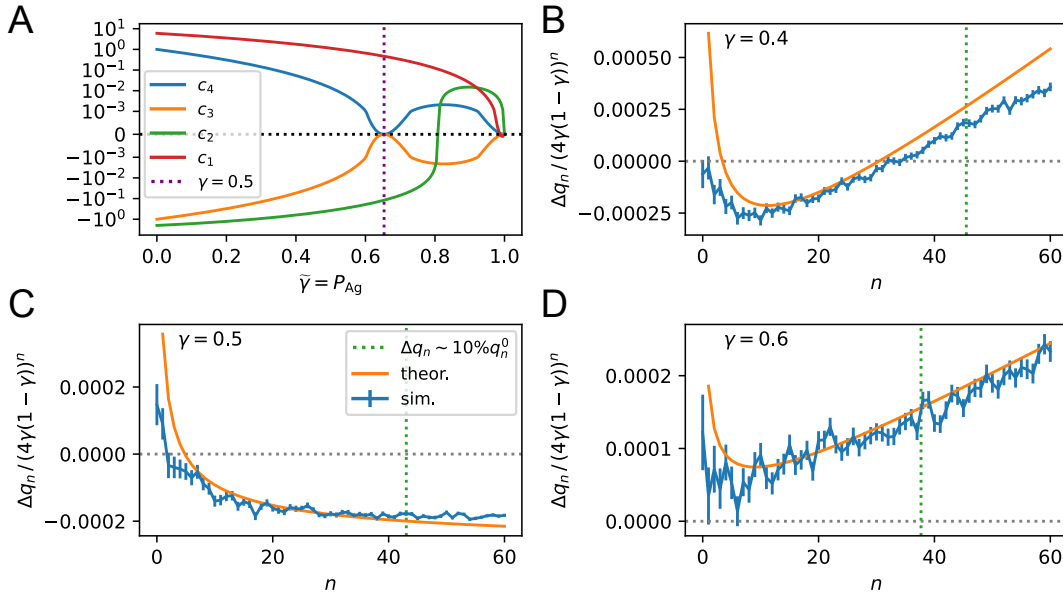


FIG. 9. **A**: values of the power expansion coefficients in eq. (43) as functions of the progenitor survival probability $\tilde{\gamma} = P_{\text{Ag}}$. Both c_4 and c_3 are null at the critical value $\gamma = 1/2$. **B,C,D**: perturbations to the probabilities q_n due to weak mutations (cf. eq. (36), with $\sigma = 0.05$) for three different progenitor affinities, corresponding to $\gamma = 0.4, 0.5, 0.6$ (B,C and D respectively). The perturbation $\Delta q_n = q_n - q_n^0$ is evaluated as the difference between the probability of lineage extinction at progeny size n without (q_n^0) and with (q_n) mutations. These were evaluated by performing 10^8 numeric simulations for each of the two conditions. In the plots we report the value $\Delta q_n (4\gamma(1-\gamma))^{-n}$ from simulations (blue) and compare it with the theoretical prediction (orange). The value of n at which $|\Delta q_n| = 10\% |q_n^0|$ sets an upper limit for the validity of the theory (green).

increase the survival probability and extinction time of low-affinity cells. Another interesting direction would be the extension of these results to the case of mutant Ags. Understanding how maturation plays out in the presence of multiple Ag mutants is currently an open issue, whose solution could lead to the development of vaccination strategies against mutable pathogens such as HIV [8, 29, 30]. One could extend our model by considering that a single cell possesses different affinities for each Ag mutant. These affinities are potentially correlated, depending on the similarity between the different mutants, and so is the effect of mutations. Taking this into account one could define mutation and selection probabilities in the presence of multiple Ag mutants, and use a similar approach to the one introduced in this paper to evaluate the lineage and population survival probability.

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Appendix A: Model parameters choice

The values of the parameters are reported in table I, and were chosen based on existing literature.

Mature GCs have a B-cells population consisting of a few thousands cell [31–33]. We therefore set the ini-

tial and maximum size of the population equal to $N_i = N_{\text{max}} = 2500$. This is in agreement with [5] which reports around 3000 cells per GC. However we stress that GCs are heterogeneous in size [25]. Similarly to [13, 19] we consider the duration of a turn of evolution to be $T_{\text{round}} \sim 12\text{h}$, which is consistent with timing of cell migration [4, 34]. Time in our model will be rescaled by this standard quantity, so that the variable t has no dimension. Similarly, also the binding energy ϵ is dimensionless, expressed in standard units of $k_B T$. For simplicity following experiments that indicate a cell-cycle time of 12h or longer [35] we consider a single division per round. Other experiments indicate an average of two cell division in the DZ [36]. We point out that, at the expense of simplicity, our theory can also be extended to account for an higher number of cell divisions.

In [18, 20] mutations occur at a rate of 0.5 per sequence per division, and are silent, lethal or affinity affecting with probabilities of respectively 0.5, 0.3, 0.2. This fixes our effective mutation probabilities to $p_{\text{sil}} = 0.75, p_{\text{let}} = 0.15, p_{\text{aa}} = 0.1$. To reproduce the fact that most of the mutations are deleterious we pick for simplicity $\mu_M = \sigma_M$. This fixes the amount of beneficial mutations to $\sim 16\%$, which is somewhat higher but still compatible with other models [13, 18, 19] in which this fraction is set to 5%. Moreover we set $\mu_M = 0.3$ so as to set the mean effect of beneficial mutations to $\langle \Delta \epsilon \rangle_{\text{beneficial}} \sim -0.15$. This value is slightly smaller than $\langle \Delta \epsilon \rangle_{\text{beneficial}} \sim -0.53$ used in [13], but this is compensated by the higher rate of

parameter	value	description
μ_i, σ_i	4, 1.5	mean and variance of the population initial binding energy distribution
N_i	2500	initial population size
N_{\max}	2500	maximum carrying capacity
$p_{\text{sil}}, p_{\text{let}}, p_{\text{aa}}$	0.75, 0.15, 0.1	probabilities of silent, lethal, affinity-affecting mutations
μ_M, σ_M	0.3, 0.3	mean and variance of distribution of affinity-affecting mutations
ϵ_{Ag}	0	Ag-binding selection threshold energy
C	see figure captions	Ag concentration
p_{diff}	0.1	differentiation probability

TABLE I. standard values of model parameters. Unless otherwise specified these are the values used in simulations. The choice of their value is discussed in appendix A

beneficial mutations in our model. The binding energy distribution of the initial population is set to a Gaussian with standard deviation $\sigma_i = 1.5$, which is compatible with experimental data [13]. Since evolution of the population is invariant for shifts of the energy space we set $\epsilon_{\text{Ag}} = 0$. Under this choice of gauge the zero in the energy space is the threshold energy for Ag-binding selection. Moreover we pick $\mu_i = 4$ so that the difference $\mu_i - \epsilon_{\text{Ag}} - \log C \sim 2\sigma_i$ for the values of Ag concentrations considered in this work ($C \sim 5$) and on average only around 2% of cells from the initial population meet this threshold. For simplicity we independently extract the energy of each cell in the initial population from this initial distribution. By doing so we might overestimate the diversity of the initial population. In fact, experiments probing the clonal composition of GCs estimated that early GCs contain around 50 to 200 different clonal families [33]. A more realistic initiation of our GCs would require us to extract the energies of around a hundred different founder cells, and let them duplicate without mutating up to to the full GC capacity. This would generate a less homogeneous initial population than the one we consider in our simulation, but would not otherwise strongly impact our results. Lastly, as in [13, 19], the probability of differentiation is set to $p_{\text{diff}} = 0.1$.

Appendix B: Estimation of $\bar{\epsilon}$ evolution

Including the effect of competition in our evaluation of the population survival probability requires us to estimate the evolution of $\bar{\epsilon}$, defined in eq. (3). To obtain a tractable approximation, we consider the deterministic limit of big population size. In this limit the population binding energy can be approximated with a continuous distribution, and the state of the system is completely determined by the density function $\rho_t(\epsilon)$. This function represents the density of cells having energy ϵ at evolution round t , so that its integral is equal to the size of the population, and its normalized version is the population binding energy distribution. Evolution is expressed in terms of operators acting on this function. In particular:

1. Cell duplication corresponds simply to doubling in

size:

$$\mathbf{A}[\rho](\epsilon) = 2 \times \rho(\epsilon) \quad (\text{B1})$$

2. Mutations are represented as the convolution with the mutation kernel $K(\Delta\epsilon)$ defined in eq. (1). Notice that the kernel K is not normalized, to account for the contribution of lethal mutations. It acts on the distribution as:

$$\mathbf{M}[\rho](\epsilon) = \int d\Delta\epsilon \rho(\epsilon - \Delta\epsilon) K(\Delta\epsilon) \quad (\text{B2})$$

3. Ag-binding selection is implemented by in the product of the population function with the survival probability eq. (2):

$$\mathbf{S}_{\text{Ag}}[\rho](\epsilon) = P_{\text{Ag}}(\epsilon) \rho(\epsilon) \quad (\text{B3})$$

4. Similarly, T-cell help selection is given by the product with the survival probability eq. (3):

$$\mathbf{S}_{\text{T}}[\rho](\epsilon) = P_{\text{T}}(\epsilon, \bar{\epsilon}) \rho(\epsilon), \quad \text{with } e^{-\bar{\epsilon}} = \frac{1}{N} \int d\epsilon e^{-\epsilon} \rho(\epsilon) \quad (\text{B4})$$

Where $N = \int d\epsilon \rho(\epsilon)$ is the current population size.

5. Differentiation consists simply in a product involving the differentiation probability:

$$\mathbf{D}[\rho](\epsilon) = (1 - p_{\text{diff}}) \rho(\epsilon) \quad (\text{B5})$$

6. Finally, the carrying capacity constraint is implemented again by a product and is operated only if the size of the population exceeds the maximum limit:

$$\mathbf{C}[\rho](\epsilon) = \min\{1, N_{\max}/N\} \rho(\epsilon) \quad (\text{B6})$$

Where again $N = \int d\epsilon \rho(\epsilon)$ is the current population size.

From these definitions the evolution of the population density function $\rho_t(\epsilon)$ can be expressed as:

$$\rho_{t+1} = \mathbf{C} \mathbf{D} \mathbf{S}_{\text{T}} \mathbf{S}_{\text{Ag}} \mathbf{M} \mathbf{A} \rho_t \quad (\text{B7})$$

Combining with the definition for $\bar{\epsilon}$ eq. (B4) provides a way for us to estimate the average evolution of $\bar{\epsilon}_t$.³

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- ³ Notice that from the order of the operators in the evolution round it follows that ϵ_t must be evaluated using eq. (B4) not directly on ρ_t but rather on $\mathbf{S}_{Ag} \mathbf{M} \mathbf{A} \rho_t$

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