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1	Is there a trade-off between horn growth and survival in adult female					
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18 Summary

19 Life history theory predicts trade-offs in energy allocation between different life history traits 20 when resources are limited, i.e. certain traits should be negatively correlated. However, 21 individuals differ in their ability to acquire resources, which can lead to positive correlations 22 between traits at the population level. Here, we investigated the consequences of the 23 allocation in horn growth and body mass on survival in a bovid (Rupicapra rupicapra) with 24 capture-mark-resighting data on 161 females. In female ungulates, body mass often covaries 25 positively with demographic performance and the few studies on horn size suggest that this trait could be a signal of individual quality. Thus, we expected to measure positive correlations 26 between the allocation in these traits and female survival. However, body mass was not 27 28 correlated to female survival and there was only a negative, though marginal, effect of horn 29 growth. Hence, it seems that the allocation in growth is not an indicator of female quality. Future studies could investigate the importance of growth on female reproduction to evaluate 30 31 its effect on lifetime reproductive success. Moreover, it is important to confirm in other 32 populations our result that suggests a cost of the allocation in horn growth to better 33 understand the presence of horns in female bovids.

Keywords: horn growth, body mass, trade-off, ungulates, life-history strategies, delayed
 costs

#### 37 Introduction

38 The existence of trade-offs is a central pillar of evolutionary theory in general and life history 39 theory in particular (Roff, 2002). When resources are limited, the allocation of energy to one function (e.g., survival, reproduction or growth) reduces the amount of energy available to 40 41 another, resulting in an observable cost. For example, reproductive costs have been described 42 in many taxa as a negative correlation between reproductive allocation and survival (Roff, 43 2002). Similarly, but less studied, we may expect a negative correlation between growth rates 44 and survival (Mangel & Stamps, 2001). However, especially in the wild, it is not always straightforward to detect trade-offs because of individual heterogeneity. Indeed, high-quality 45 individuals may be able to acquire more resources and may thus allocate more energy, 46 47 simultaneously, in several traits than low-quality individuals (van Noordwijk & de Jong, 1986; 48 Wilson & Nussey, 2010). This results in positive correlations at the population level even if an 49 underlying trade-off exists at the individual level (Hamel et al., 2009) (individual quality 50 hypothesis). Moreover, the costs of an allocation may only be expressed at later stages: for 51 example higher growth rates early in life can be correlated with reduced later performance and shorter lifespans (Metcalfe & Monaghan, 2003). 52

53 Here, we investigated the correlations between the allocation in growth and survival in 54 females of a long-lived mammal, the chamois (Rupicapra rupicapra). We used two variables 55 to describe growth: body mass and horn length. Body mass often covaries positively with demographic performance in ungulates, for example it can be positively correlated with 56 57 juvenile survival (Côté & Festa-Bianchet, 2001) and negatively with female age at primiparity 58 (Green & Rothstein, 1991). Thus, it seems to be an indicator of female quality and we expect 59 it to covary positively with survival under the 'individual quality hypothesis'. However, this 60 does not preclude the existence of a delayed cost of growth, which would be revealed by a

61 negative correlation in old females. In bovids, such as chamois, horns are permanent and both 62 males and females wear horns. While male horns are known for their role in sexual selection (male-male competition and female choice), the role and putative associated costs of horns 63 are much less studied in females. We aim here to clarify the correlation between horn length 64 and age-specific survival in a female bovid. Given that large-horned females may be dominant 65 66 over small-horned females (Locati & Lovari, 1991) and attain primiparity at an earlier age 67 (Rughetti & Festa-Bianchet, 2011), horn size may be a measure of individual quality and thus 68 may covary positively with survival (i.e., support for the individual quality hypothesis). However, horn growth may be energetically costly (e.g. it is adjusted to food availability 69 (Chirichella et al., 2013 and references therein)), and may represent a source of heat loss 70 71 during winter (Picard et al., 1994). Therefore, large-horned females may bear the cost of their 72 larger horns later in life, in which case we predict a negative covariation between horn growth 73 and old-female survival.

74 Material and methods

75 Site and studied species

76 A population of chamois (Rupicapra rupicapra, Linnaeus, 1758), a mountain ungulate, has 77 been monitored since the early 1980s in the Game and Wildlife Reserve (GWR) of the Bauges 78 massif (45°40'N, 6°13'E, 5,205 ha), in the northern French Alps. This site has been described 79 elsewhere (Loison, Jullien, & Menaut, 1999). Chamois are trapped below falling nets baited 80 with salt (June to September). Measurements on captured individuals included body mass, 81 total horn length and horn growth from birth to 3 years of age (we averaged the measures 82 from both horns). Recaptures are rare in this population and thus we used only the measures 83 at first capture. Age was determined by counting horn annuli (Schröder & von Elsner-Schack, 84 1985). Animals were individually marked with a coloured collar and visually monitored between April and December each year with a varying intensity of observation pressure per year (see supplementary materials). For this study, we considered data from the site Armenaz for females  $\geq$  3 years of age for which early horn growth (0-3 year increment) was measured. The GWR is a protected area where only guided hunt occurs. We are therefore confident that all marked animals that were shot were reported (n = 20, early horn growth = 162.78 ± 16.14 mm). The data were right-censored (i.e. females were removed from the analyses after being shot) to estimate survivorship independently of any hunting effect (Lebreton *et al.*, 1992).

#### 92 <u>Statistical analyses</u>

93 Body mass varies with age and capture date. In all the analyses, we used a corrected measure 94 of body mass defined as the residuals of a linear regression of log(body mass) against capture 95 date (F<sub>1,154</sub>=15.41, P<0.001) and age (linear: F<sub>1,154</sub>=11.26, P<0.001, quadratic: F<sub>1,154</sub>=6.31, 96 P=0.013). The interactions between age (linear and quadratic terms) and capture date were 97 not included (all P>0.70). For early horn growth, we used the standardized (i.e. centered and 98 reduced) early horn growth and performed a Pearsons's correlation with standardized total 99 horn length and one with corrected body mass. Total horn length was standardized within 100 each age (except for females aged 10 and more that were pooled).

101 We tested whether age-specific survival was correlated to early horn growth (0-3 year 102 increment) or body mass at capture, using these 2 measurements as individual covariates in 103 Capture-Mark-Recapture modelling in E-SURGE (version 1.9.0) (Choquet, Rouan, & Pradel, 104 2009b). We first tested the goodness-of-fit of the Cormack-Jolly-Seber model using U-CARE 105 (version 2.3.2) (Choquet et al., 2009a). The Cormack-Jolly-Seber model satisfactorily fitted the 106 data (chi = 63.70, P = 0.382). There was no transience (Z = -2.11, P = 0.982), but a strong "trap-107 happiness" (Z = -5.26, P < 0.001, see examples of trap-dependence lato sensu , i.e. when 108 animals are not physically recaptured, in (Pradel & Sanz-Aguilar, 2012)). This effect was well

109 accounted for with models taking into account individual heterogeneity in behaviour (e.g. 110 some animals temporarily changing their home range; see supplementary materials). From 111 the model with heterogeneity, we selected the best model describing the time effect in re-112 sighting probabilities (Table S1). Then, we fitted models on the survival probabilities. We 113 proceeded in two steps. First, we tested for the importance of age and time (year effect). We 114 considered 3 typical age classes in ungulates: 3-7 years old, 8-12 years old and >12 years old (Loison, Gaillard, & Houssin, 1994; Gaillard et al., 2004). Second, we tested for effects of 115 116 individual covariates: in the best model selected in the first step, we added an effect of early 117 horn growth (standardized measure) or body mass (corrected measure) on survival of all the age classes or only one age class. Model selection was based on AICc: the model with the 118 119 strongest support is the one with the lowest AICc. It is considered that two models differ when 120 their difference of AICc is higher than 2 (Burnham & Anderson, 1998). We also calculated the 121 AICc weight of each model (Burnham, Anderson, & Huyvaert, 2011).

122 **Results** 

123 We gathered data on 161 females captured between 1992 and 2012 aged 3 to 12 years old 124 (Table S2). Horn growth between 0 and 3 averaged 160.7 mm (SE = 14.0 mm, Figure S1) and 125 accounted for most of total horn length (Figure 1). There was a positive correlation between 126 early horn growth and total horn length corrected for age (R = 0.61, t = 9.62, df = 158, P < 127 0.001, Figure S2). At capture, female body mass averaged 29.23 kg (SE = 3.23kg, Figure S3). 128 Early horn growth and body mass corrected for age and capture date were positively 129 correlated (R = 0.18, t = 2.27, df = 153, P = 0.025, Figure S4). Survival probabilities varied among 130 years and age-classes (Table 1). Once these effects were accounted for, the best-fitting model 131 included a negative effect of early horn growth on yearly survival of 8-12 years old females 132 (Figure 2, model M4 in Table 2). The model with a negative effect on all females also had a

similar AICc (model M2 in Table 2). However, the baseline model showed the same support in
the data than these two models (Model M0 in Table 2). The best-fitting model M4 is only 1.40
times more likely than model M0. There was no effect of body mass on survival (Table 2).
Model M4 is 2.9 times more likely than model M8 which is the model with the lowest AICc
within models including an effect of body mass.

138 **Discussion** 

139 We investigated the correlation between early horn growth, body mass and survival in adult 140 female chamois. Interestingly, early horn growth and adult body mass were positively 141 correlated, which suggests that these variables can be a measure of some of the same aspects of female growth. In this population, horn length at age 3 accounted on average for 85% of 142 143 horn length when older than 8 years old. This confirms some previous results that showed 144 that most horn growth occurs during the first summers of female life in chamois (Bassano, 145 Perrone, & Von Hardenberg, 2003; Rughetti & Festa-Bianchet, 2011). Contrary to our 146 predictions and other findings in ungulates (e.g. Gaillard et al., 2000; Poissant et al., 2008), 147 adult body mass was not correlated with annual age-specific survival. We measured body 148 mass only once and the fact that body mass does not seem to be a relevant predictor of annual 149 survival may reflect the inter- and intra-annual variability of this variable. Concerning horn 150 size, the model with most support in the data showed a delayed cost of growing large horns: 151 old females with longer horns tended to have a lower survival than old females with shorter 152 horns. The absence of a highly significant effect may be due to a lack of power because of the 153 sample size and/or because of the large annual variation in survival between years. Indeed, 154 trade-offs are more easily detected in years with overall low survival. Further studies will be 155 needed to better understand the influence of annual variations on horn growth and on the 156 trade-off with survival. Nevertheless, it is surprising to detect a cost, because even in male

157 ungulates with larger horns than chamois, there is no clear pattern linking horn size and 158 survival (e.g., see table 1 of Bonenfant et al., 2009; Toïgo, Gaillard, & Loison, 2013). Further 159 studies in other populations and other species should clarify the existence of such cost in female bovids. For example, there is no correlation between longevity and horn volume in 160 female bighorn sheep (Poissant et al., 2008). If our result is confirmed in chamois and females 161 162 bear a cost of wearing large horns, this suggests that large horns may be counter-selected in 163 females. It would be particularly interesting to know whether the same pattern is observed in 164 males or if selection for this trait is acting in opposite direction in both sexes. However, the presence of horns in females may also be important in other aspects of female life and 165 performance, such as in sociality and reproduction (Locati & Lovari, 1991; Rughetti & Festa-166 167 Bianchet, 2011). Accordingly, females with different growth patterns may differ in other life-168 history traits, such as reproductive effort. In this case, our result may indicate a reproductive cost. To clarify this aspect, future studies will need to investigate the covariation between 169 170 female phenotype and reproductive trajectory (Hamel *et al.*, 2009, 2010).

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#### 178 **References**

Bassano B, Perrone A, Von Hardenberg A. 2003. Body weight and horn development in
 Alpine chamois, *Rupicapra rupicapra* (Bovidae, Caprinae). *Mammalia* 67: 65–74.

- 181 **Bonenfant C, Pelletier F, Garel M, Bergeron P**. **2009**. Age-dependent relationship between 182 horn growth and survival in wild sheep. *Journal of Animal Ecology* **78**: 161–171.
- Burnham KP, Anderson DR. 1998. Model selection and inference: a practical information theoretical approach. New-York, NY: Springer-Verlag.
- 185 Burnham KP, Anderson DR, Huyvaert KP. 2011. AIC model selection and multimodel
- 186 inference in behavioral ecology: some background, observations, and comparisons.
- 187 Behavioral Ecology and Sociobiology **65**: 23–35.
- 188 Chirichella R, Ciuti S, Grignolio S, Rocca M, Apollonio M. 2013. The role of geological
   189 substrate for horn growth in ungulates: a case study on Alpine chamois. *Evolutionary Ecology* 190 27: 145–163.
- Choquet R, Lebreton JD, Gimenez O, Reboulet AM, Pradel R. 2009a. U-CARE: Utilities for
   performing goodness of fit tests and manipulating CApture-REcapture data. *Ecography* 32:
   1071–1074.
- 194 **Choquet R, Rouan L, Pradel R. 2009b**. Program E-Surge: a software application for fitting
- 195 multievent models. In: Thomson DL, Cooch EG, Conroy MJ, eds. Environmental and
- Ecological Statistics. Modeling Demographic Processes In Marked Populations. Springer US,845–865.
- Côté SD, Festa-Bianchet M. 2001. Birthdate, mass and survival in mountain goat kids: effects
   of maternal characteristics and forage quality. *Oecologia* 127: 230–238.
- Gaillard JM, Festa-Bianchet M, Delorme D, Jorgenson J. 2000. Body mass and individual
   fitness in female ungulates: bigger is not always better. *Proceedings of the Royal Society of London. Series B: Biological Sciences* 267: 471–477.
- Gaillard JM, Viallefont A, Loison A, Festa-Bianchet M. 2004. Assessing senescence patterns
   in populations of large mammals. *Animal Biodiversity and Conservation* 27: 47–58.
- Green WCH, Rothstein A. 1991. Trade-offs between growth and reproduction in female
   bison. *Oecologia* 86: 521–527.
- Hamel S, Côté SD, Gaillard JM, Festa-Bianchet M. 2009. Individual variation in reproductive
   costs of reproduction: high-quality females always do better. *Journal of Animal Ecology* 78:
   143–151.
- 210 Hamel S, Gaillard JM, Yoccoz NG, Loison A, Bonenfant C, Descamps S. 2010. Fitness costs of
- 211 reproduction depend on life speed: empirical evidence from mammalian populations.
- 212 *Ecology letters* **13**: 915–935.
- 213 Lebreton JD, Burnham KP, Clobert J, Anderson DR. 1992. Modeling survival and testing
- biological hypotheses using marked animals: a unified approach with case studies. *Ecological*
- 215 *Monographs* **62**: 67–118.
- 216 Locati M, Lovari S. 1991. Clues for dominance in female chamois: age, weight, or horn size?
- 217 *Aggressive Behavior* **17**: 11–15.

- Loison A, Gaillard JM, Houssin H. 1994. New insight on survivorship of female chamois
- 219 (*Rupicapra rupicapra*) from observation of marked animals. *Canadian Journal of Zoology* 72:
   220 591–597.
- Loison A, Jullien JM, Menaut P. 1999. Subpopulation structure and dispersal in two populations of chamois. *Journal of Mammalogy* **80**: 620.
- Mangel M, Stamps J. 2001. Trade-offs between growth and mortality and the maintenance of individual variation in growth. *Evolutionary Ecology Research* **3**: 583–593.
- Metcalfe NB, Monaghan P. 2003. Growth versus lifespan: perspectives from evolutionary
   ecology. *Experimental Gerontology* 38: 935–940.
- Van Noordwijk AJ, de Jong G. 1986. Acquisition and allocation of resources: their influence
   on variation in life history tactics. *The American Naturalist* 128: 137–142.
- Picard K, Thomas DW, Festabianchet M, Lanthier C. 1994. Bovid horns an important site
   for heat-loss during winter. *Journal of Mammalogy* 75: 710–713.
- Poissant J, Wilson AJ, Festa-Bianchet M, Hogg JT, Coltman DW. 2008. Quantitative genetics and sex-specific selection on sexually dimorphic traits in bighorn sheep. *Proceedings of the*
- 233 Royal Society B: Biological Sciences **275**: 623–628.
- Pradel R, Sanz-Aguilar A. 2012. Modeling trap-awareness and related phenomena in
   capture-recapture studies. *PLoS ONE* 7: e32666.
- 236 **Roff DA**. **2002**. *Life history evolution*. Sunderland, MA: Sinauer Associates.
- Rughetti M, Festa-Bianchet M. 2011. Effects of early horn growth on reproduction and
   hunting mortality in female chamois. *Journal of Animal Ecology* 80: 438–447.
- Schröder W, von Elsner-Schack I. 1985. Correct age determination in chamois. In: Lovari S,
   ed. The Biology and management of mountain ungulates. London, UK: Croom Helm, 65–70.
- Toïgo C, Gaillard JM, Loison A. 2013. Alpine ibex males grow large horns at no survival cost
   for most of their lifetime. *Oecologia* 173: 1261–1269.
- Wilson AJ, Nussey DH. 2010. What is individual quality? An evolutionary perspective. *Trends in Ecology & Evolution* 25: 207–214.
- 245
- 246

#### 247 Figure captions

- Figure 1. Relationship between early horn growth (0-3 year horn increment) and total horn
- length for each age at capture. Early horn growth does not include the growth that occurs
- after the 3<sup>rd</sup> winter and thus is not equal to total horn length of individuals aged 3 years old.

251

- 252 **Figure 2.** Relationship between early horn growth (0-3 year horn increment) and survival for
- 253 females of the 8-12 age class. The lines represent the average estimates and average
- confidence intervals for all years of the model M4 presented in Table 2. The distribution of
- 255 early horn growth in our population is at the bottom of the figure.

- 257 **Tables**
- 258 **Table 1.** Effects of age class and year on female survival modelled with capture-mark-

recapture data. The estimates of age-specific survival from the "A1 + A2 + A3 + year" model

are shown as the average of the estimates for each year with the average confidence

261 intervals.

Model	k	Deviance	AICc	ΔAICc	
A1 + A2 + A3 + year	32	1070.74	1138.05	0	
year	30	1086.73	1149.64	11.59	
A1 + A2 + A3	12	1121.04	1145.52	7.47	
Constant survival	10	1149.58	1169.92	31.87	
	A1	A2	А	A3	
Mean estimates	0.94	0.88 (		).79	
	[0.75 ; 0.99]	[0.63 ; 0.96] [0.50		; 0.93]	
Moon offect size (legit scale)	2.84	1.97	1.	1.33	
	[1.09 ; 4.37]	[0.54 ; 3.31]	[0.01	; 2.61]	

262 A1 = 3-7 years old

263 A2 = 8-12 years old

A3 = > 12 years old

266 **Table 2.** Effects of horn growth and body growth on female survival modelled with capture-

267 mark-recapture data. The models with the strongest support in the data are in bold (ΔAICc <

268 2 and AICc weight > 0.10). The values of the effect size are on the logit scale.

	Model	k	Deviance	AICc	ΔAICc	AICc	Effect
U						weight	size
MO	base = A1 + A2 + A3 + year	32	1070.74	1138.05	0.67	0.172	
M1	base + HS(A1) + HS(A2) + HS(A3)	35	1067.75	1141.73	4.35	0.027	
M2	base + HS	33	1068.85	1138.38	1.00	0.146	-0.22 ± 0.16
M3	base + HS(A1)	33	1070.64	1140.16	2.78	0.060	
M4	base + HS(A2)	33	1067.85	1137.38	0	0.241	-0.43 ± 0.27
M5	base + HS(A3)	33	1070.56	1140.09	2.71	0.062	
M6	base + BM(A1) + BM(A2) + BM(A3)	35	1069.51	1143.49	6.11	0.011	
M7	base + BM	33	1070.36	1139.89	2.51	0.069	
M8	base + BM(A1)	33	1069.99	1139.52	2.14	0.083	
M9	base + BM(A2)	33	1070.29	1139.82	2.44	0.071	
M10	base + BM(A3)	33	1070.69	1140.22	2.84	0.058	

269 BM = corrected body mass (corrected for age and date of capture)

270 HS = standardized early horn growth (horn increment between 0-3 years)

271 A1 = 3-7 years old

A2 = 8-12 years old

273 A3 = > 12 years old

HS(A1) = effect of horn size on survival of the individuals of the first age class