

Modelling European small pelagic fish distribution: Methodological insights

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Alexandre Schickele, Boris Leroy, Gregory Beaugrand, Eric Goberville, Tarek Hattab, et al.. Modelling European small pelagic fish distribution: Methodological insights. Ecological Modelling, 2020, 416 (1), pp.212-225. 10.1016/j.ecolmodel.2019.108902. hal-02514810

HAL Id: hal-02514810 https://hal.science/hal-02514810v1

Submitted on 6 Jan 2021

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1	<u>Title:</u> Modelling European small pelagic fish distribution: methodological insights
2	
3	Authors: Alexandre Schickele ^{1*} , Boris Leroy ² , Gregory Beaugrand ^{3,4} , Eric Goberville ² , Tarek
4	Hattab ⁵ , Patrice Francour ^{1§} & Virginie Raybaud ^{1§}
5	
6	[§] Both authors contributed equally to this work.
7	*Corresponding author
8	¹ Université Côte d'Azur, CNRS, UMR 7035 ECOSEAS, Nice, France.
9	² Unité Biologie des Organismes et Ecosystèmes Aquatiques (BOREA), Muséum National
10	d'Histoire Naturelle, Sorbonne Université, Université de Caen Normandie, Université des
11	Antilles, CNRS, IRD, Paris, France.
12	³ CNRS, Univ. Lille, Univ. Littoral Côte d'Opale, UMR 8187, LOG, Laboratoire d'Océanologie
13	et de Géosciences, Wimereux, France.
14	⁴ Sir Alister Hardy Foundation for Ocean Science, The Laboratory, Citadel Hill, Plymouth,
15	United Kingdom.
16	⁵ MARBEC, Univ. Montpellier, CNRS, Ifremer, IRD, Avenue Jean Monnet, Sète, France
17	
18	Authors' additional information:
19	Alexandre Schickele: alex.schickele@gmail.com (ORCID ID: 0000-0002-4842-6890)
20	Boris Leroy: boris.leroy@mnhn.fr (ORCID ID : 0000-0002-7686-4302)
21	Gregory Beaugrand: gregory.beaugrand@univ-lille1.fr (ORCID ID: 0000-0002-0712-5223)
22	Eric Goberville: eric.goberville@upmc.fr (ORCID ID: 0000-0002-1843-7855)
23	Tarek Hattab: tarek.hattab@ifremer.fr (ORCID ID: 0000-0002-1420-5758)
24	Patrice Francour: patrice.francour@univ-cotedazur.fr (ORCID ID: 0000-0002-8489-1319)
25	Virginie Raybaud: virginie.raybaud@univ-cotedazur.fr (ORCID ID: 0000-0001-6426-3326)

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26 <u>ABSTRACT</u>

The distribution of marine organisms is strongly influenced by climatic gradients 27 worldwide. The ecological niche (sensu Hutchinson) of a species, i.e. the combination of 28 29 environmental tolerances and resources required by an organism, interacts with the environment 30 to determine its geographical range. This duality between niche and distribution allows climate change biologists to model potential species' distributions from past to future conditions. While 31 32 species distribution models (SDMs) have been intensively used over the last years, no 33 consensual framework to parametrise, calibrate and evaluate models has emerged. Here, to 34 model the contemporary (1990-2017) spatial distribution of seven highly harvested European 35 small pelagic fish species, we implemented a comprehensive and replicable numerical procedure based on 8 SDMs (7 from the Biomod2 framework plus the NPPEN model). This 36 37 procedure considers critical issues in species distribution modelling such as sampling bias, 38 pseudo-absence selection, model evaluation and uncertainty quantification respectively through 39 (i) an environmental filtration of observation data, (ii) a convex hull based pseudo-absence 40 selection, (iii) a multi-criteria evaluation of model outputs and (iv) an ensemble modelling 41 approach. By mitigating environmental sampling bias in observation data and by identifying the most ecologically relevant predictors, our framework helps to improve the modelling of fish 42 43 species' environmental suitability. Not only average temperature, but also temperature 44 variability appears as major factors driving small pelagic fish distribution, and areas of highest 45 environmental suitability were found along the north-western Mediterranean coasts, the Bay of Biscay and the North Sea. We demonstrate in this study that the use of appropriate data pre-46 47 processing techniques, an often-overlooked step in modelling, increase model predictive 48 performance, strengthening our confidence in the reliability of predictions.

50	Keywords: species distribution models, pseudo-absence, sampling bias, convex hull,				
51	uncertainty, small pelagic fish				
52					
53	Highlights:				
54	• Temperature (mean and variability) are major species distribution drivers				
55	• Small pelagic species have high environmental suitability along western Europe				

- Environmental filtering compensates spatial sampling bias
- The convex hull method is a robust pseudo-absence selection strategy

59 1. INTRODUCTION

60 Fish species distribution and assemblages are strongly influenced by both climatic and physical gradients (Ben Rais Lasram et al. 2010, Beaugrand et al. 2011, Raybaud et al. 2017). 61 62 Temperature is known as a master parameter driving fish distribution at a macroecological level (Lenoir et al. 2011, Beaugrand et al. 2018). This parameter influences a large range of biological 63 processes such as growth, reproduction, larval development, recruitment, and act as a major 64 stressing factor depending on species thermal tolerance (psychrophile or thermophile species; 65 66 Angilletta 2011, Beaugrand and Kirby 2018). Salinity, oxygen concentration, primary production (that are indirectly influenced by changes in temperature; e.g. Kirby and Beaugrand 67 2009) or the physical habitat (e.g. sediment type; Poloczanska et al. 2013) may also highly 68 69 influence marine fish species at different spatial scales.

70

Hutchinson (1957) conceptualised the ecological niche as the "n-dimensional ensemble 71 72 of environmental conditions that enable a species to live and reproduce" and subsequently made 73 a distinction between the fundamental and the realised niche (Hutchinson 1978). Due to biotic 74 interactions, dispersal limitation and/or historical factors (Soberon and Peterson 2005), species generally occupy only their realised niche, i.e. the subset of their fundamental niche that 75 76 represents the response of all physiological processes of a species to the synergistic effects of 77 environmental conditions (Helaouet and Beaugrand 2009, Beaugrand et al. 2013). By defining 78 the niche as an attribute of species instead of a portion of the environment, the Hutchinson's 79 concept enables duality between niche and distribution (Pulliam 2000, Colwell and Rangel 80 2009). Such a relationship is of major interest in biogeography as each georeferenced species 81 occurrence, i.e. where a given species has been observed, can be related to several 82 environmental parameters such as temperature, salinity and primary production. When species are in equilibrium with their environment, associating environmental conditions and observed 83

84 distributions permits climate change biologists to estimate species' potential niche (Jiménez85 Valverde et al. 2008).

86

87 The relationship between species occurrences, environmental conditions and species' potential niche has become intensively studied over the last two decades, using a wide range of 88 modelling techniques - hereafter referred to as Species Distribution Models (SDMs) to assess 89 past, contemporary and future species distribution in both marine and terrestrial ecosystems 90 91 (e.g. Cheung et al. 2009, Bellard et al. 2016, Cristofari et al. 2018). SDMs rely on several 92 ecological assumptions, such as species distribution in equilibrium or habitat saturation 93 (Soberon and Peterson 2005), niche conservatism (Crisp et al. 2009), unlimited dispersal 94 abilities (Wiens et al. 2009) or the non-influential role of biotic interactions in shaping largescale distributions (i.e. the Gleasonian vision of biotic communities; Gleason 1926, Guisan and 95 96 Thuiller 2005, Wiens et al. 2009). Superimposed to these assumptions, several sources of errors 97 and uncertainties may lead to variation - sometimes conflicting - in the outputs of SDMs for a 98 given species (Beaumont et al. 2008): (i) accuracy of observation data and (ii) lack of true 99 absences (Proosdij et al. 2016), (iii) identification of ecologically meaningful environmental 100 predictors with high explanatory power (Guisan and Thuiller 2005), (iv) choice of the modelling 101 algorithm (Buisson et al. 2010) and (v) SDMs' evaluation processes (Leroy et al. 2018). While 102 tremendous progresses have been made on both the building and evaluation of SDMs in recent 103 years with a plethora of new methods for modelling species' distribution (Araújo and Guisan 104 2006, Leroy et al. 2018, Støa et al. 2018), the development of further procedures is still required 105 for improving the quality of SDMs.

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Species distribution models are known to be very sensitive to different sources ofuncertainties and sustained attention should be devoted to each step of the modelling procedure,

109 from the pre-processing of species occurrences data to model evaluation. Such an approach is 110 essential to increase confidence in model outputs (Porfirio et al. 2014): for most areas of the 111 world and species, survey effort often exhibits strong spatial and temporal bias, occurrence 112 records being frequently too scarce, constrained to presence-only data or both. Working with 113 biased observation datasets may result in under- or over-estimated species distributional ranges 114 (Araújo and Guisan 2006, Dormann et al. 2007), leading therefore to inaccurate modelled 115 contemporary distributions, which are inadequate for assessing potential future range shifts or 116 for defining conservation measures. Similarly, biased pseudo-absence datasets (e.g. multiple 117 pseudo-absences selected in the same environmental conditions or coinciding with 118 environmental conditions where the species is observed) may lead to a distorted estimation of 119 species distributional ranges (e.g. Wisz and Guisan 2009, Lobo and Tognelli 2011). A 120 modelling framework that includes a preliminary stage devoted to the construction of a 121 representative calibration dataset – as well as its associated level of uncertainty assessment – is 122 therefore essential (e.g. Varela et al. 2014).

123

Here, we developed a framework that encompasses recent advances on the building, calibration and evaluation of SDMs with the aim of (i) selecting the most relevant environmental parameters, (ii) generating consistent pseudo-absence data and (iii) validating representative model outputs (Cornwell et al. 2004, Varela et al. 2014, Leroy et al. 2018).

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We applied this framework on seven economically important European Small Pelagic
Fish (SPF) species (Mediterranean horse mackerel *Trachurus mediterraneus*, Atlantic horse
mackerel *Trachurus trachurus*, European pilchard *Sardina pilchardus*, round sardinella *Sardinella aurita*, European sprat *Sprattus sprattus*, European Anchovy *Engraulis encrasicolus*and bogue *Boops boops*). These seven SPF species are widely distributed planktonic feeders

known for their central role in marine food webs (Cury 2000, Checkley et al. 2009). Moreover,
they are of major economic importance and represent a large part of the Mediterranean and
Black Sea commercial landings (more than 50% between 2000 and 2013; FAO 2016).
However, while SPFs are ideal candidates for SDMs because of their sensitivity to
environmental factors (Perry et al. 2005), their European distribution is far from being
exhaustively documented and available records originated from diverse and/or nonstandardised monitoring surveys (FAO 2016).

141

142 2. MATERIAL AND METHODS

143 <u>2.1. Biological and environmental data</u>

144 2.1.1. Small pelagic fish occurrence data

145 Occurrence records (e.g. fisheries independent trawl surveys, discrete research 146 samplings) for the seven SPF species (Mediterranean horse mackerel, Atlantic horse mackerel, 147 European pilchard, round sardinella, European sprat, European Anchovy and Bogue) were 148 compiled from three available public databases: the Ocean Biogeographic Information System 149 Mapper (OBIS, http://www.iobis.org/mapper/), the Global Biodiversity Information Facility 150 (GBIF, https://www.gbif.org/) and Fishbase (http://www.fishbase.org/). When possible, we 151 included observations retrieved from the literature to construct the most up-to-date datasets 152 encompassing their entire distribution range (see Supplementary material 1). Biological data 153 retrieved for our study ranged from 1950 to 2017, recent records (since 1990) prevailing 154 (83.2±6.7 %) over both past (1950-1990; 12.2±8.7 %) and undated observations (4.6±3.6 %). 155 Past or undated records were only considered along the distribution edge when the species 156 presence was confirmed by recent records. This precautionary approach avoided over- or under-157 predictions of the model due to low quality presence data (Kramer-Schadt et al. 2013). The observation records pre-processing consisted in a data cleaning procedure applied on each 158

species dataset to (i) remove unreliable observations (e.g. preserved specimen; Newbold 2010) 159 160 and false identifications (e.g. taxonomic confusion), (ii) discard duplicate occurrences and (iii) 161 ensure the temporal and locational reliability at the edge of the observed distribution (e.g. data 162 on land, longitudinal and/or latitudinal inversion, historical or undated data). According to the 163 ecology of SPFs - species cannot be observed below 300 m depth (Checkley et al. 2009) -164 while remaining permissive, a precautionary bathymetry threshold (-1000 m) was applied to 165 remove inconsistent occurrences. Following this pre-processing, we obtained seven clean 166 datasets, with a number of observations ranging from 1314 (for Mediterranean horse mackerel) 167 to 24806 (for European sprat). For the seven SPFs, occurrences were aggregated on a 0.1° x 168 0.1° spatial grid (from 70°N to 70°S and 180°E to 180°W) that corresponds to that of 169 environmental parameters.

170

171 2.1.2 Environmental data

To calculate the ecological niche (sensu Hutchinson, 1957) of each SPF, we collected 172 173 environmental parameters from different databases (see Table 1 for details). Environmental 174 parameters values for each spatial grid cell were first calculated for each year and then averaged 175 on the 1990-2017 period. The environmental parameters presented in **Table 1** were retrieved in 176 different spatial resolutions ranging from 0.1° to 0.5°. For modelling purpose, all variables were 177 therefore interpolated to a $0.1^{\circ} \times 0.1^{\circ}$ grid using a bilinear interpolation in the geographical domain available for all environmental parameters, ranging from 70°N to 70°S and 180°E to 178 179 180°W.

180

181 2.2. Description of the models

We used two approaches to model the potential environmental suitability (i.e. spatialised
index between 0 and 1, defined as a probability of presence based on environmental parameters)

of each SPF species over the 1990-2017 period: (i) the Non-Parametric Probabilistic Ecological 184 185 Niche (NPPEN; Beaugrand et al. 2011) model and (ii) seven modelling algorithms available 186 within the BIOMOD2 package (Thuiller et al. 2016). The NPPEN model is a presence only 187 model based on the Mahalanobis generalised distance (Mahalanobis 1936) and on a modified 188 version of the Multiple Response Permutation Procedure (MRPP; Mielke et al. 1981). The 189 BIOMOD2 framework allows ensemble modelling of species distribution (i.e. an average 190 model of a wide range of algorithms; Thuiller et al. 2009). Here, seven algorithms were 191 considered: (i) Generalised Linear Model (GLM), (ii) Generalised Additive Model (GAM), (iii) 192 Generalised Boosting Model (GBM), (iv) Artificial Neural Network (ANN), (v) Flexible 193 Discriminant Analysis (FDA), (vi) Multiple Adaptive Regression Splines (MARS) and (vii) 194 Random Forest (RF). Because the models used in this study have been already described and 195 discussed in several publications (e.g. Beaugrand et al. 2011, Lenoir et al. 2011, Raybaud et al. 196 2015 for NPPEN, e.g. Thuiller et al. 2009, Albouy et al. 2012, Bellard et al. 2013 for 197 BIOMOD2), we refer the reader to this literature for further information. The algorithms were 198 calibrated using the default parameters in BIOMOD2, optimised for species distribution 199 modelling (details in Thuiller et al. 2016). By including this large range of algorithms within 200 an ensemble model approach, we quantified the uncertainty related to the selection of SDMs 201 (Pearson et al. 2006, Buisson et al. 2010) by calculating the standard deviation (SD) and the 202 coefficient of variation (CV) among SDM outputs.

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204 <u>2.3. Data preparation and ensemble model selection</u>

205 2.3.1. Pre-selection of the environmental parameters and assessment of multicollinearity

To model the ecological niche of the seven SPFs, we first constructed the full dataset of environmental parameters based on our knowledge of the ecology of SPFs. A variable selection process (**Figure 1**, step 1) was then applied to identify, at the species level, the most 209 parsimonious dataset that explained each species distribution. This process follows the 210 procedure described in Leroy et al. (2014) and Bellard et al. (2016). Because most of the 211 algorithms (especially regression-based models such as GLM) are sensitive to multicollinearity - that may distort model estimation (Dormann et al. 2013) - relations among environmental 212 213 parameters were assessed by means of the Pearson correlation coefficient, using a threshold r > 214 0.7 to reduce the initial environmental matrix. When two or more environmental parameters 215 showed correlation values above this threshold, only one variable was retained (details in 216 Supplementary material 2).

217

218 We subsequently assessed the relative importance of each environmental parameter by 219 sequentially randomising each variable and by calculating the resulting current distribution 220 (Leroy et al. 2014). The variables that best predicted SPF distribution were sea surface 221 temperature annual mean (SST), temperature variability (sea surface temperature annual range 222 or monthly variance, depending on the targeted species), bathymetry and distance to coast (see 223 Supplementary material 2). In order to avoid model over-parameterisation (that affects model 224 performance, model transferability and assessment of variable importance), we chose not to 225 include bathymetry and distance to coast directly in the models, but in a hierarchical filtering 226 approach (Hattab et al. 2014): for a given geographical cell, environmental conditions were 227 considered as suitable for a marine species only if a probability of occurrence coincided with a 228 distance to coast less than 50km or up to a 300m depth for oceanic cells, i.e. outside the 50km 229 wide coastal area. Concerning environmental predictors, we systematically considered 230 temperature (mean and variability) in our models. Finally, we tested the relevance of including 231 sea surface salinity (SSS) and/or primary production (log_PP) as a potential third explanatory 232 environmental parameter in the models. Each run is detailed in Supplementary material 3.

234 2.3.2. Environmental filtration and pseudo-absence selection

Because sampling effort is neither homogeneous and nor standardised over marine regions, occurrence data may not be representative of the whole populations, a requirement to increase the reliability of SDMs (Lobo and Tognelli 2011). While under-sampling is commonly observed at the edge of species range (Varela et al. 2014), observation datasets can also be biased toward regions more comprehensively investigated due to an easy access or a long tradition of monitoring (Fithian et al. 2015).

241 To consider the risk of over-sampling, and the ensuing over-representation of 242 environmental features (Kramer-Schadt et al. 2013), we first homogenised species datasets to 243 assign the same weight to over- and under-sampled regions (Figure 1, step 2). A 244 multidimensional matrix was designed for each species and each combination of environmental 245 parameters, a dimension reflecting an environmental factor. Each cell of the homogenised 246 matrix was considered as an environmental stratum, i.e. a combination of a set of parameters, 247 with the following resolution: 0.5°C for temperature-related parameters, 0.5 for SSS and 0.5 mol.m⁻².s⁻¹ (in log) for primary production. In case an environmental stratum contained multiple 248 249 occurrences, only one occurrence (i.e. one 0.1° x 0.1° geographical cell with the corresponding 250 environmental conditions) was kept in the homogenised dataset.

251 We also considered the lack of absence data. To assess this gap, we generated pseudo-252 absences using the convex hull method (Cornwell et al. 2004, Getz and Wilmers 2006). The 253 convex hull was defined here as the smallest convex hyper-volume in the environmental space 254 containing all species observation records. A restricted convex hull (see Figure 2) has been 255 defined as a convex hull excluding occurrence points within the 2.5% and 97.5% percentiles 256 for each environmental parameter (i.e. excluding observations in the most extreme 257 environmental conditions). This restricted convex hull is considered as a proxy of the suitable 258 environmental conditions outside which, pseudo-absences were randomly generated in equal 259 number to the filtered occurrences as recommended by the "D-designs" theory (Montgomery 260 2005): the optimal design to minimise prediction variance is when an equal number of 261 observations are at opposite value extremes (Montgomery 2005, Hengl et al. 2009) and when 262 there is a high spreading in the feature space. Finally, for each species, pseudo-absence were 263 projected back in geographical cells showing environmental conditions outside SPF species' environmentally favourable areas (Figure 2; Varela et al. 2014). Finally, model outputs 264 265 obtained from our environmental filtration approach were compared with outputs for which 266 neither environmental filtration nor the convex hull pseudo-absence selection method was 267 applied (Figure 3).

268

269 2.3.3. Validation and selection of the best models

270 We then quantified the performance of our models using five commonly used evaluation 271 metrics: (i) the Continuous Boyce Index (CBI; Hirzel et al. 2006), a metric specifically designed 272 for presence-only models and insensitive to pseudo-absences, (ii) the Area Under the Curve 273 (AUC; Swets 1988, Fielding and Bell 1997), (iii) the True Skill Statistic (TSS; Allouche et al. 274 2006), (iv) the Jaccard and (v) the Sørensen similarity indices (Jaccard 1908, Sørensen 1948). However, because all evaluation metrics – except the CBI – require both presence and absence 275 276 data (see discussion in Leroy et al. 2018 about the use of pseudo-absence to evaluate the 277 performance of models) and because some may be affected by prevalence (i.e. the ratio between 278 the number of observed presence and generated pseudo-absence; Leroy et al. 2018) we based 279 our selection process of the best models on CBI values only. We considered models to be wrong when CBI values were below -0.5, "average to random" for values ranging from -0.5 to 0.5, 280 281 and good for values above 0.5 (Faillettaz et al. 2019).

For each model, we computed evaluation metrics by performing a cross-validation procedure with 10 repetitions. We randomly sampled 70% of the occurrence data to calibrate

the model and kept the remaining 30% for model validation (Merow et al. 2013). Following the 284 285 "evaluation strip method" detailed by Elith et al. (2005), the adequacy between observed and 286 modelled spatial distributions was also assessed by means of response curves. For a given 287 environmental parameter, the corresponding response curve was calculated, while keeping the 288 other parameters constant (i.e. at the mean value corresponding to their occurrence points). By 289 doing this, we identified spurious results (e.g. we do not expect bimodal responses to 290 temperature) and/or unexpected distribution ranges (e.g. large portions of predicted range in 291 regions where the species has never been observed, and vice-versa; Supplementary material 4).

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- 293

3. RESULTS

295 <u>3.1. SDMs and parameters selected in the ensemble models</u>

Based on the calculation of the CBI values and the examination of species response curves (Supplementary material 3 and 4), we identified the best models for each SPF species. Our results showed that both GLM and NPPEN models were almost always selected in the ensemble model, except for the European anchovy.

Ensemble models showed that temperature-related variables were essential to assess the 300 301 spatial distribution of SPFs'. For virtually all species, the models that considered mean 302 temperature and variability showed high ability to reproduce the overall SPFs distributions 303 (Table 2, Supplementary material 3) with CBI values always above 0.5 (Faillettaz et al. 2019). 304 However, some discrepancies were observed among species. While Mediterranean horse 305 mackerel, Atlantic horse mackerel and European anchovy distributions were more related to 306 mean monthly temperature variance (SSTvar), European pilchard, round sardinella, European 307 Sprat and bogue distributions were better reproduced when mean annual temperature range (SSTr) was considered. Despite the high correlation between SSTr and SSTvar (r=0.80, 308

Supplementary material 2), both variables have dissimilar ecological influences (seasonality versus short-term climatic variability respectively). Primary production also emerged as important to model species' spatial distribution. Finally, we highlighted the important role of sea surface salinity (SSS) for both European pilchard and European anchovy, by discriminating both the Baltic and the Black seas from other regions (**Table 2**).

314

315 By applying our environmental filtration framework, we substantially improved the 316 modelling of most of the SPFs spatial distributions (Figure 3, individual contributions of the 317 filtration process and the convexhull are presented in Supplementary material 5), except for the 318 European pilchard (Figure 3b). Specifically, we observed an increase in mean CBI values that 319 ranged from +0.05 to +0.23 (Figure 3). For most SPFs, lower Environmental Suitability Index 320 (ESI) values were obtained (-0.2 without filtration to -0.6 with filtration), suggesting that our 321 procedure alleviated the risk of over-prediction, especially in the Black and Baltic seas, and 322 beyond 60°N where species have never been observed (Figure 4, left panels). By increasing 323 ESI values from +0.4 to +0.6, environmental filtration also emphasised regions known to be 324 highly suitable for SPF species, but in which occurrences were only scarcely available (e.g. in 325 the eastern Mediterranean Sea for Atlantic horse mackerel, round sardinella and bogue; Figure 326 4a, f and g).

327 <u>3.2. Contemporary (1990-2017) environmental suitability of small pelagic fishes</u>

We then represented the contemporary (1990-2017) spatial distribution of the seven SPFs in the spatial domain ranging from 10 to 70°N and from 30°W to 45°E (**Figure 4**, middle panel) Environmental suitabilities at the calibration range (i.e. the entire distribution range) are provided in Supplementary material 6.

According to the observed and modelled distributions (Figure 4, left and middle 333 334 panels), two species groups were identified with respect to their environmental suitability along 335 the European coasts. The first group encompassed temperate-to-cold water species (hereafter 336 "temperate-cold" species; i.e. Atlantic horse mackerel, European pilchard, European sprat and 337 European anchovy; Figure 4a-d) that were more likely to be present in northern Europe. The 338 second grouped temperate-to-warm water species (hereafter "temperate-warm" species; i.e. 339 Mediterranean horse mackerel, round sardinella and bogue; Figure 4e-g) located along the 340 Mediterranean coasts down, to North Africa.

341 The four temperate-cold species showed the highest ESI values in the North Sea, in the 342 Celtic Sea, in the Bay of Biscay (ESI values > 0.8) and to a lesser extent along Norwegian 343 coasts (ESI values ranging from 0.2 to 0.8). For all temperate-cold species, but European 344 pilchard, high ESI values (from 0.4 to 0.8) were expected in the western and central regions of 345 the Baltic Sea (Figure 4), suggesting that these species can tolerate a wide salinity range (from 346 8 to 38) and a high thermal variability (up to 20°C annual range). All temperate-cold species, 347 but European sprat, showed high ESI values (from 0.6 to 0.8) in the north-western part of the 348 Mediterranean basin (Figure 4). For all temperate-cold species, the modelled ESIs are in 349 accordance with the observation data except in southern Iceland, western Norway and to a lesser 350 extent in the eastern Black Sea where positive ESI values (between 0.05 to 0.6) are predicted 351 while no observed distribution is available.

The three temperate-warm species showed the highest ESI values (from 0.4 to 0.8) in nearly all the regions of the Mediterranean Sea and medium to low ESI values (from 0.2 to 0.7) in the Black Sea and along the north-western African coasts. However, some discrepancies among species were detected (**Figure 4**). Round sardinella appears as the most southern SPF species with no suitable conditions north of the Portuguese coast. On the contrary, Mediterranean horse mackerel and bogue showed high ESI values (from 0.6 to 0.8) along the

358 Atlantic coasts from the Celtic sea down to northern Africa, up to 0.8 in the Bay of Biscay. 359 While bogue showed maximum ESI values (> 0.8) in the whole Mediterranean Sea, only the 360 north-western regions of the Mediterranean Sea were highly suitable for Mediterranean horse 361 mackerel and round sardinella. The modelled ESIs are in accordance with the observation data 362 except in the North Sea for Mediterranean horse mackerel and Bogue and to a lesser extent in the eastern Black Sea for all temperate-warm species. These regions highlight positive ESI 363 364 values (between 0.05 and 0.6) while no observations are available. These discrepancies may 365 result from an absence of sampling in these regions or external factors hindering species 366 establishment despite suitable environmental conditions.

367

368 <u>3.3. Model uncertainties</u>

369 Two main sources of uncertainties in projected species distributions were considered in 370 our study: (i) biological uncertainties, related to the quality of occurrence datasets and (ii) 371 numerical uncertainties, inherent to the selection of different modelling algorithms (Pearson et 372 al. 2006, Buisson et al. 2010). Standard deviations (SD) - computed, for each species, from 373 outputs that originated from both selected algorithms and cross-validation runs – ranged from 374 0.1 to 0.4, indicating a convergence between models (Figure 4, right panels). The lowest SD 375 values (close to 0.2) were found in the north-western Mediterranean Sea for virtually all SPFs, 376 and in the Bay of Biscay and in the North Sea when temperate-cold species were studied 377 (Figure 4, a-d). The highest SD values (close to 0.4) were observed in the Mediterranean Sea 378 for Mediterranean horse mackerel, European pilchard and round sardinella (Figure 4, e-g). For all species, the coefficient of variation (CV; Supplementary material 7) highlighted very low 379 380 CV variations (< 20%) towards their centre of distribution (in the Mediterranean Sea for all 381 species and North Sea for temperate-cold species) while showing high variations at the leading 382 or the trailing edge of their distribution (up to 100% in the Black, Baltic and the Norwegian383 seas).

384

385 4. DISCUSSION

386 By combining several numerical techniques such as the convex hull method, the 387 ensemble models approach and an examination of species response curves in a comprehensive 388 modelling framework, we modelled the contemporary (1990-2017) environmental suitability 389 of seven of the most commercially and ecologically important European small pelagic fish. By 390 relying on both an understanding of the ecological requirements of species and on the use of 391 innovative statistical tools, our framework allowed us to focus only on the best models, to 392 improve the way species distribution modelling is carried out, and therefore to produce more 393 robust ecological scenarios.

394

395 At a macroecological level, thermal-induced effects have been frequently related to 396 latitudinal mean temperature gradients (Angilletta 2011). While our analysis showed that mean 397 temperature (SST) had a major influence on species distributions, we also revealed the key role 398 of temperature seasonality (SSTr) and short-term temperature variations (SSTvar) in shaping 399 distributional ranges (Table 2). Small pelagic fishes are marine ectotherms, that mainly depend 400 on external heat sources, their body temperature being directly controlled by environmental 401 conditions directly (Checkley et al. 2009). Changes in temperature may therefore affect SPFs' 402 physiological performances (i.e. their fitness; Perry et al. 2005, Payne et al. 2016). Because the 403 relationship between temperature and fitness occurred through species' thermal optimum and 404 range, and because SPFs are short lifespan species (Checkley et al. 2009), annual temperature 405 changes may affect several life stages (especially reproduction and larval development; e.g. 406 Peck et al. 2013), with long-term consequences on population dynamics (Fréon et al. 2005).

407 Small pelagic fishes may also experience ontogenetic shifts in thermal tolerance during their 408 development (Peck et al. 2013) and temperature seasonality (here SSTr) may either favour or 409 perturb species development, with potential consequences on distributional patterns (Figure 4, 410 middle panels; Peck et al. 2013). This is especially noticeable in regions characterised by an 411 important thermal variability, such as in the Black and Azov seas, in the Northern Adriatic Sea, 412 in the Baltic Sea and to a lesser extent in the eastern part of the North Sea. Considering thermal variability in SDMs (e.g. the monthly SST variance) may therefore help to better define species 413 414 environmental suitability and to minimise the risk of over-prediction at the leading and the 415 trailing edges of their distributions (Lenoir et al. 2011).

416

417 When used in distribution modelling, regression-based algorithms such as GLM, are 418 known to be rather sensitive to environmental sampling bias, which may induce type I errors 419 (i.e. false positive), with consequences on projected species environmental suitability (Araújo 420 and Guisan 2006, Dormann et al. 2007). However, as for many other species (e.g. Boakes et al. 421 2010), commonly available databases of SPFs provide a distorted view of their actual 422 distribution because of spatial and temporal bias in species observations (e.g. Beck et al. 2014). 423 When the time comes to evaluate the quality of biodiversity datasets, three major issues have 424 been raised in the literature (e.g. Kramer-Schadt et al. 2013, Guillera-Arroita et al. 2015): the 425 influence of (i) prevalence, i.e. the proportion of sites in which the species was recorded as 426 present, (ii) imperfect species detection and (iii) sampling bias. Despite an increasing 427 availability of information, the biogeographic distribution of most species remain still 428 frequently incomplete (Bini et al. 2006); a shortcoming explained, inter alia, by heterogeneous 429 sampling effort among surveys, or the inaccessibility of some areas. For all SPF datasets, this 430 effect is undeniable when comparing the north-western Mediterranean Sea, the Bay of Biscay, the North Sea with other European regions. (Figure 4, left panels). To lower this issue, a 431

plethora of data sources (e.g. standardised scientific surveys, biodiversity portals) are now 432 433 available in collaborative databases (e.g. GBIF), offering more cohesive summaries of species' 434 distributions although leading - sometimes - to enhanced spatial and environmental biases 435 (Kramer-Schadt et al. 2013, Beck et al. 2014). Considering independent distributional data (i.e. 436 from private collections or from the literature; Beck et al. 2013) along with the associated pre-437 processing (e.g. Kramer-Schadt et al. 2013, Varela et al. 2014, Aiello-Lammens et al. 2015, 438 Fithian et al. 2015), can contribute to cover the ecological niches of species more 439 comprehensively and to improve model accuracy. By coupling these procedures with our 440 restricting convex hull pseudo-absence selection, we (i) assigned the same weight to 441 environmental conditions independently of the observation density (i.e. alleviating observation 442 sampling bias), (ii) lowered the weight of presence records at the distribution edge (i.e. avoiding 443 the risk of over-prediction) and (iii) selected unbiased pseudo-absence (i.e. independent of the 444 observation bias).

445

446 Applying environmental filtering and the restricted convexhull pseudo-absence 447 selection method resulted in ensemble models characterised by a reduced ESI in over-sampled 448 areas and an increased ESI in undersampled areas. Our results are consistent with our 449 expectations and in line with previous studies that suggested that random generation of pseudo-450 absences and/or a selection process based on geographical criterion may lead to lower 451 predictability (e.g. Wisz and Guisan 2009, Hattab et al. 2014). Although real absences lead to 452 higher model accuracy (Wisz and Guisan 2009), they are rarely available (Boakes et al. 2010) 453 and determining the location of pseudo-absences on the basis of a statistical analysis such as 454 the convex hull is a reliable alternative (Hattab et al. 2013). Finally, our approach limits 455 spurious species response curves (i.e. overfitted or bimodal curves; Supplementary material 4) 456 and decreases the risk of over-predictions towards the edge of the species range. We

457 acknowledge that we may have slightly underpredicted the European pilchard distribution in 458 Kattegat (i.e. strait between Denmark and Sweden); the high amount of occurrence records 459 slightly outside the modelled distribution in this region may have biased the calculation of the 460 CBI. Despite the well-known robustness of this index (Breiner et al. 2015, Faillettaz et al. 2019), 461 our result highlight that no evaluation metric is optimal and that both comparison between 462 observed and modelled distributions and examination of species responses curves are essential 463 for assessing the reliability of model outputs.

464

465 While the assessment of the environmental suitability for a given species may differ -466 slightly or markedly – from one SDM to another (Pearson et al. 2006, Buisson et al. 2010), it is 467 still challenging to identify the most appropriate model (see discussion in Buisson et al. 2010). Even if several methods have been recently proposed, no consensus has emerged (see 468 469 discussion in Leroy et al. 2018). and the use of different - well-fitted and evaluated - SDMs 470 may help to better simulate potential species distributions, for past, contemporary and future 471 environmental conditions (Araújo and New 2007). In complementarity with a multi-SDM 472 approach, we think that researchers should examine species response curves during the 473 evaluation process (e.g. Elith and Leathwick 2009, Jarnevich et al. 2018, Erauskin-Extramiana 474 et al. 2019). As observed for Mediterranean horse mackerel (see details in Supplementary 475 material 4), we invalidated response curves that were statistically significant but not in 476 agreement with the ecological niche theory. Without this complementary evaluation method, 477 the corresponding algorithms would have been considered in the ensemble model, therefore 478 potentially resulting in spurious patterns of ESIs. Therefore, this multi-criteria evaluation 479 procedure is of great interest from a (i) numerical (i.e. metric adapted to presence-only datasets) 480 and an ecological (i.e. validation of the species-environment relationships) perspective. Note 481 that the seven SPFs we chose are representative of a large spectrum of environmental

482 conditions, from temperate-to-cold waters (e.g. European sprat) to temperate-to-warm waters
483 (e.g. bogue and round sardinella). To conclude, our framework has been faced with a wide range
484 of environmental conditions, allowing us to better evaluate its robustness, sensitivity and
485 possible transferability to other species and ecosystems.

486

487 In this work, we have estimated species' potential niche and not the realised niche 488 (Soberón and Nakamura 2009). We caution that additional environmental parameters, 489 biological interactions and species life traits (e.g. dispersal abilities) may explain why we 490 detected environmentally suitable conditions in regions where SPFs were not observed (e.g. the 491 Norwegian Sea; Pulliam 2000, Pearman et al. 2008). Considering the role of biotic interactions 492 in shaping species distributions (Chaalali et al. 2016) would improve the reliability of SDMs 493 outputs by better estimating and simulating the realised niche of species (Wisz et al. 2013, 494 Louthan et al. 2015). Including dispersal mechanisms while accounting for oceanic currents 495 and physical barriers after the potential distribution modelling step may help to refine the 496 distributional range of species (Engler and Guisan 2009). These approaches require an 497 exhaustive ecological understanding of the interaction process at a macroecological scale and a 498 deep knowledge of species life traits to implement metrics that simulate the ability of species 499 to disperse (e.g. Petitgas et al. 2012). Moreover, it is important to notice that no direct 500 correlations between ESI (potential or realised) and spatialised biomass or official catches have 501 been established in the literature although temporal correlations have been suggested however 502 (e.g. Chaalali et al. 2016). Therefore, discrepancies between SPF's ESI, biomass and official 503 catches (e.g. FAO 2016) may be explained by population-related parameters (e.g. recruitment, 504 growth, biotic interaction) or management policies and stock status (e.g. under or over-fishing), 505 respectively. Finally, inter-specific absolute ESI comparison is challenging because of the 506 monospecific nature of SDMs.

508 Our study presents a detailed environmental suitability assessment of seven of the most 509 heavily harvested European SPFs. By focusing on the most common sources of errors and 510 uncertainties in SDMs, we designed a comprehensive - fully transferable to other species and 511 ecosystems - modelling framework which is intended to elaborate more robust ecological 512 scenarios. Our framework addressed several critical steps in SDMs, i.e. the treatment of 513 sampling biases in observation records, the generation relevant pseudo-absences and a dual 514 assessment of model outputs that proposes to evaluate models from both a numerical and an 515 ecological perspective. In a conservation decision-making perspective, these different steps are 516 essential to increase confidence in SDMs, a prerequisite to propose effective resource 517 management measures (e.g. accounting for environmental stress) or to measure the 518 effectiveness of protected areas (e.g. regarding environmental resilience). Moreover, when used in combination with scenarios of future environmental conditions (i.e. IPCC climate scenarios), 519 520 this framework provides robust contemporary predictions to assess possible changes in species 521 distribution in the context of global climate change. Despite the growing literature on the 522 development and testing of new modelling and evaluation processes, the use of SDMs in 523 quantitative resource management and scientific surveys is still a great challenge.

524

525 DECLARATIONS

Acknowledgements - This paper is dedicated to the memory of Prof. Patrice Francour who passed away on October 13th, 2019. He devoted his life to the protection and understanding of Mediterranean ecosystems, initiated and co-supervised this work. We thank J.O. Irisson for his support on the NPPEN R code. We also acknowledge the Ocean Biogeographic Information System (OBIS), the Global Biodiversity Information Facility (GBIF) and Fishbase for providing species observation data. Finally, we thank the IPCC Coupled Model

532	Intercomparison Project (CMIP) and the climate modelling groups for making available their
533	model output.

Funding - This work was supported by the Prince Albert II of Monaco foundation through the

project CLIM-ECO². AS's PhD is funded by the Provence-Alpes-Côte-d'Azur (PACA) Region

537	in partnership with the Comité Régional des Pêches Maritimes et des Elevages Marins
538	(CRPMEM) de PACA.
539	
540	Author contribution - VR and PF conceived and supervised the study. AS, VR and EG collected
541	the data. AS performed the numerical analysis. BL, GB, TH, EG and VR helped in the
542	modelling process. AS and EG wrote the first draft. BL provided the code to use BIOMOD2.
543	All authors made substantial contributions in the successive versions of the manuscript.
544	
545	Conflicts of interest – All authors have no conflict of interest to declare
546	
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774 TABLES AND FIGURES

Name (Period)	Description	Reference	
Bathymetry	Spatial seafloor depth (m)	Global seafloor topography (Smith and Sandwell 1997)	
Distance to coast	Distance to the nearest coast (km)	NASA Goddard Space Flight Center (2009) (https://oceancolor.gsfc.nasa.gov/docs/distfromcoa st/)	
SSS (1990-2017)	Sea Surface Salinity	Levitus' climatology (Levitus 2011) completed with (http://www.ices.dk/)	
PP (1990-2017)	Sea Surface Primary Production (mol.m ⁻² .s ⁻¹). Averaged from five general circulation models (IPSL, MPI, CNRM, HadGEM and GISS).	IPSL (Dufresne et al. 2013, Hourdin et al. 2013), MPI (Stevens et al. 2013, Giorgetta et al. 2013), CNRM (Voldoire et al. 2013), HadGEM (Jones et al. 2011)	
Log_PP (1990-2017)	Log10-transformed Sea Surface Primary Production	and GISS (Schmidt et al. 2014) models.	
SST (1990-2017)	Mean annual Sea Surface Temperature (°C)	AVHRR Very High Resolution Radiometer (Casey et al. 2010)	
SSTmax (1990-2017)	Mean sea surface temperature of the hottest month (°C)		
SSTmin (1990-2017)	Mean sea surface temperature of the coldest month (°C)		
SSTr (1990-2017)	Mean annual sea surface temperature range (°C). Difference between SSTmax and SSTmin.		
SSTvar (1990-2017)	Mean monthly sea surface temperature variance (°C). Calculated using monthly SST data.		

Table 1: Environmental parameters used to model SPF distribution.

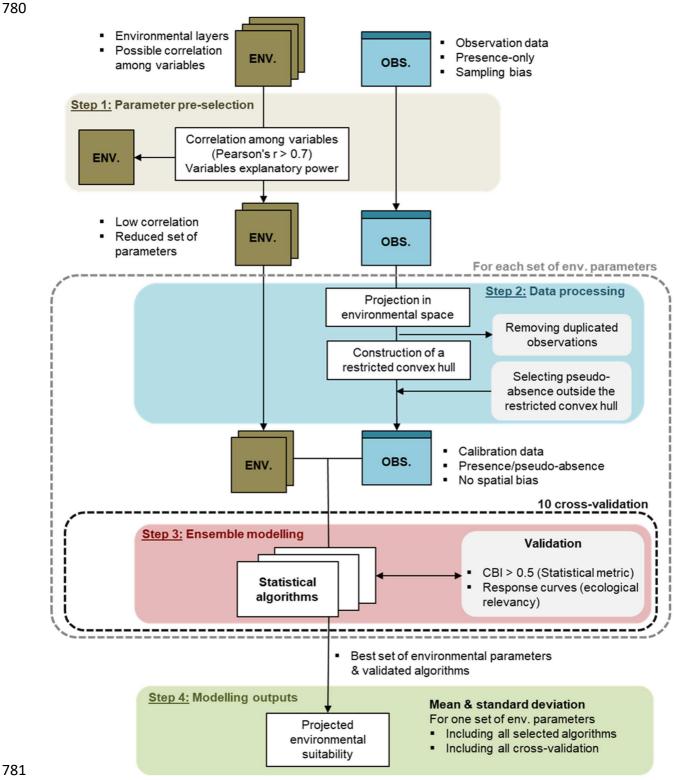


Figure 1: Sketch diagram of the modelling framework applied to model SPFs species. "ENV." = environmental



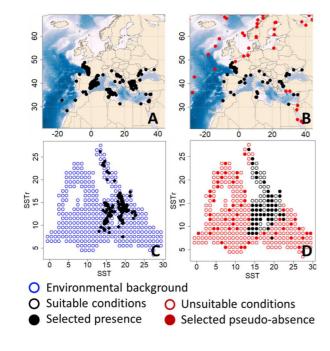


Figure 2: Example of pseudo-absences generation for the Mediterranean horse mackerel (environmental parameters = SST + SSTr, 1°C resolution). A-C: Species occurrences (black dots) in (A) the geographical domain and (C) the environmental space. B-D: Species occurrences (black dots) and pseudo-absences (red dots) generated from the restricted convex hull method in (B) the geographical domain and (D) the environmental space.

Mediterranean horse mackerel	Parameters:	SST, SSTvar, log_PP
	Models:	GLM, RF, NPPEN
	CBI (mean):	0.71
Atlantic horse mackerel	Parameters:	SST, SSTvar, log_PP
	Models:	GLM, NPPEN
	CBI (mean):	0.95
European pilchard	Parameters:	SST, SSTr, SSS
	Models:	GLM, GAM, NPPEN
	CBI (mean):	0.75
Round sardinella	Parameters:	SST, SSTr, log_PP
	Models:	GLM, RF, FDA, NPPEN
	CBI (mean):	0.88
European sprat	Parameters:	SST, SSTr, log_PP
	Models:	GLM, MARS, NPPEN
	CBI (mean):	0.92
European anchovy	Parameters:	SST, SSTvar, SSS
	Models:	GLM, FDA, MARS
	CBI (mean):	0.88
Bogue	Parameters:	SST, SSTr
	Models:	GLM, ANN, NPPEN
	CBI (mean):	0.65

Table 2: Environmental parameters and SDMs selected by our procedure.

The selected SDMs had a CBI>0.5 and satisfying response curves. Parameters: (SST) Sea Surface Temperature,
(SSTr) annual range of Sea Surface Temperature, (SSTvar) monthly variance of Sea Surface Temperature,
(log_PP) log-transformed Primary Production and (SSS) Sea Surface Salinity. Models: (GLM) Generalised
Linear Model, (GAM) Generalised Additive Model, (GBM) Generalised Boosting Model, (ANN) Artificial
Neural Network, (FDA) Flexible Discriminant Analysis, (MARS) Multiple Adaptive Regression Splines, (RF)
Random Forest and (NPPEN) Non Parametric Probabilistic Ecological Niche model.



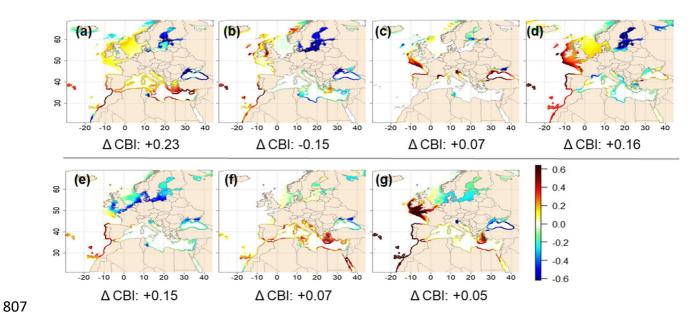


Figure 3: Environmental suitability index and CBI differences between ensemble models originating from our
modelling framework and ensemble models constructed without data filtration and random pseudo-absence
selection for (a) Atlantic horse mackerel, (b) European pilchard, (c) European sprat, (d) European anchovy, (e)
Mediterranean horse mackerel, (f) round sardinella and (g) bogue.

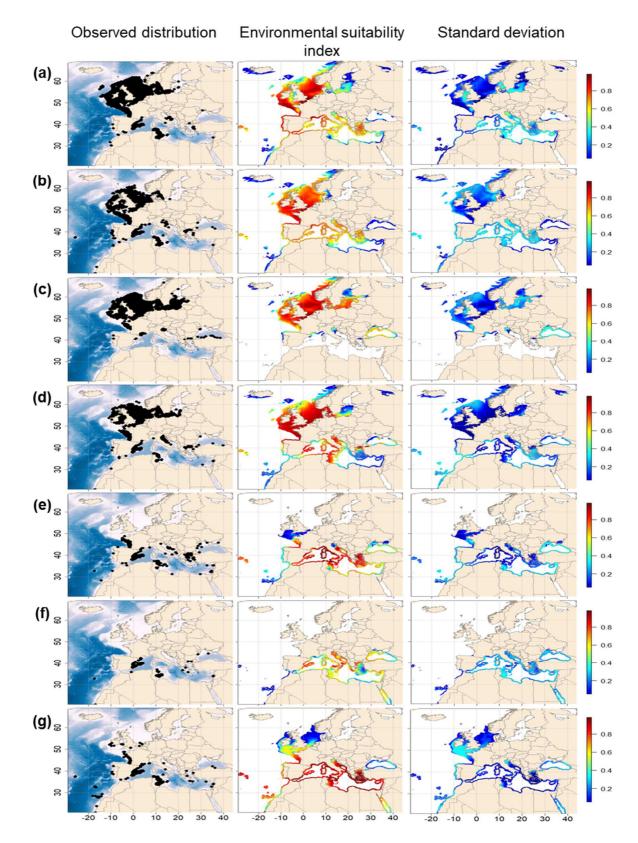


Figure 4: Contemporary (1990-2017) observed distribution (left panels), modelled environmental suitability index
(0 to 1, middle panels) and its associated standard deviation (0 to 1, based on all validated SDMs and crossvalidation runs, right panels) for (a) Atlantic horse mackerel, (b) European pilchard, (c) European sprat, (d)
European anchovy, (e) Mediterranean horse mackerel, (f) round sardinella and (g) bogue.