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► **To cite this version:**

| Claudiu Tiberiu Albuлесcu. Coronavirus and oil price crash. 2020. hal-02507184v1

**HAL Id: hal-02507184**

**<https://hal.science/hal-02507184v1>**

Preprint submitted on 12 Mar 2020 (v1), last revised 18 Mar 2020 (v2)

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# Coronavirus and oil price crash: A note

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## Abstract

Coronavirus (COVID-19) creates fear and uncertainty, hitting the global economy and amplifying the financial markets volatility. The oil price reaction to COVID-19 was gradually accommodated until March 09, 2020, when, 49 days after the release of the first coronavirus monitoring report by the World Health Organization (WHO), Saudi Arabia floods the market with oil. As a result, international prices drop with more than 20% in one single day. Against this background, the purpose of this paper is to investigate the impact of COVID-19 numbers on crude oil prices, while controlling for the impact of financial volatility and the United States (US) economic policy uncertainty. Our ARDL estimation shows that the COVID-19 daily reported cases of new infections have a marginal negative impact on the crude oil prices in the long run. Nevertheless, by amplifying the financial markets volatility, COVID-19 also has an indirect effect on the recent dynamics of crude oil prices.

**Keywords:** oil price; coronavirus; COVID-19; financial volatility; economic policy uncertainty

**JEL codes:** Q41, G15, G41

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## 1. Introduction

The coronavirus (COVID-19) pandemic outbreaks Wuhan (Hubei region from China), where the first infection case is reported on December 31, 2019. 49 days after the release of the first coronavirus monitoring report by the World Health Organization (WHO) on January 21, 2020, over 100,000 people from more than 100 countries around the world are affected. Although COVID-19 does not present similar patterns in terms of fatality rate as compared to the 2002-2003 Severe Acute Respiratory Syndrome (SARS) (3.5% for COVID-19 and 11% for SARS), or in terms of global spread as compared to the Spanish Flu pandemic of 1919 (which affected hundreds of millions of people), it is considered very contagious and triggers a lot of uncertainty in the real economy and financial markets. COVID-19 negatively affects the overall demand, creating short-run volatility in the food prices,<sup>1</sup> and impeding the mobility of workers and tourists. In addition, COVID-19 creates fear and additional stress on financial markets, where the price volatility is continuously increasing. Anticipating a strong decrease in the global demand in the next period, Saudi Arabia starts an oil price war on March 09, 2020, and floods the market with oil. In one single day the crude oil price plunges with more than 20%. This shock spills over financial markets that crash during the same day (the Black Monday).

In this context, the present note tempts to investigate how the COVID-19 new infection cases affect the oil price, while controlling for the role of financial stress and volatility (VIX index), and the United States (US) economic policy uncertainty (EPU). The relationship between stock prices and oil prices was intensively debated in the literature given the financialization of commodity markets (for a recent review, see [Balcilar et al., 2019](#)). However, the interaction between financial volatility on the one hand, and oil prices on the other hand, receives less attention. [Illing and Liu \(2006\)](#) notice that oil price shocks correspond to picks in the financial stress index for Canada, and formally test this relationship. In fact, financial stress episodes affect the stock market returns and create volatility, whereas the stock prices and oil prices are highly correlated. At the same time, oil price shocks generate financial volatility and uncertainty. This bidirectional relationship is documented by [Nazlioglu et al. \(2015\)](#), who conduct a volatility spillover analysis on the US, for the period 1991 to 2014. The authors report the dominance of the long-run volatility spillover. More recently, [Das et al. \(2018\)](#) examine the dependence between stock prices,

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<sup>1</sup> Supermarkets were emptied in several regions from Italy, Germany and the United Kingdom.

commodity prices, and financial stress. The authors resort to a nonparametric causality-in-quantile technique and focus on the US economy. They document a bilateral causality between oil price and financial volatility. The result confirms the findings by [Basher and Sadorsky \(2016\)](#) who investigate the hedging property of oil prices and VIX (the implied financial volatility of S&P 500), for emerging markets stock prices.

Another strand of literature investigates the relationship between oil prices and economic uncertainty, resorting to the EPU index constructed by [Baker et al. \(2016\)](#). This relationship is particularly important for the US, the dollar being the transaction currency on oil markets. Oil prices influence the forecasts of macroeconomic variables and therefore the EPU. At the same time, policy-induced uncertainty impacts the asset prices in general, including oil prices. In this line, [Kang et al. \(2017\)](#) examine the dynamic effects between the US EPU and international oil prices and discover that US oil production shocks are associated with a decrease of economic uncertainty. Similar findings are reported by [Antonakakis et al. \(2014\)](#) who resort to the spillover index proposed by [Diebold and Yilmaz \(2012\)](#) and report that EPU shocks negatively respond to oil price shocks and *vice-versa*. [Hailemariam et al. \(2019\)](#) extend the analysis at the level of G7 countries using monthly data from 1997 to 2018. They underline the existence of a time-varying effect of oil prices on EPU.

Other studies document opposite findings, highlighting the impact of uncertainty on oil price returns. A nonparametric causality-in-quantiles method is proposed by [Shahzad et al. \(2017\)](#) to investigate the causal effect of investors' sentiments and EPU on commodity prices, including oil. With a focus on extreme dependencies and resorting to a copula approach, [Aloui et al. \(2016\)](#) show that an increased EPU index has a positive effect on oil prices in periods that precede a financial crisis outburst. [Yang \(2019\)](#) performs both a causality and a spillover analysis between EPU and oil price shocks and emphasizes that oil prices behave as net receivers of information from EPU. A different approach is implemented by [Albulescu et al. \(2019\)](#) who show that the US EPU influences the connectedness between oil and currency markets.

Finally, a separate set of studies assesses the interaction between both economic uncertainty and financial volatility, and oil prices. In a quantile regression framework, [Reboredo and Uddin \(2016\)](#) test to what extent the financial stress and policy uncertainty impact the energy and metal commodity prices in the US. They report a nonlinear effect of EPU and VIX on oil prices. With a focus on the shocks in oil prices, [Degiannakis et al. \(2018\)](#)

document a time-varying effect on the US EPU and financial uncertainty, in specific periods.<sup>2</sup> However, none of these studies investigates the uncertainty, fear and panic triggered by COVID-19 on oil prices.

Therefore, our first contribution to the exiting literature is the assessment of the impact of COVID-19 on oil prices, while controlling for the role of US EPU and VIX. As far as we know, this is the first paper approaching this subject, which will certainly modify the public agendas for the following months. We use the WHO official announcements in terms of COVID-19 new infection cases, and the West Texas Intermediate (WTI) for crude oil prices.

Second, resorting to daily data, we look to both the short- and long-run relationship between oil prices, VIX and US EPU. Our series are either I(1) or I(0), which recommends the use of the Autoregressive Distributed Lag (ARDL) approach proposed by [Pesaran et al. \(2001\)](#). Third, we assess the effect of COVID-19 on oil prices, looking to the impact of the total number of new infection cases, as well as to the new cases reported in China and outside China. Finally, we test the robustness of our findings considering the BRENT crude prices.

The rest of the paper presents data and methodology (Section 2), the empirical findings (Section 3) and the policy implications (Section 4). Finally, we conclude.

## **2. Data and methodology**

### **2.1. Data**

COVID-19 data are obtained from the daily situation reports released by WHO starting with January 21, 2020. Consequently, our sample covers the period January 21, 2020 – March 09, 2020 (49 observations). The oil price data are obtained from the US Energy Information Administration (EIA), whereas the financial volatility (VIX) data comes from the Chicago Board Options Exchange (CBOE). The US EPU daily data are extracted from [Baker et al. \(2016\)](#) and daily updated for the US on their website.<sup>3</sup>

Figure 1 presents the dynamics of the indicators. Figure 1(a) shows that crude oil prices record a continuous decline since the official monitoring of COVID-19, and a severe crash starting with March 6, 2020. During this period the financial volatility and economic policy uncertainty were continuously growing. On February 17, 2020 the Chinese authorities

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<sup>2</sup> We notice that the relationship between economic uncertainty, financial volatility and oil prices is characterized by reverse causality. In fact, even the economic policy uncertainty represents a reliable cause of financial markets volatility as shown by [Antonakakis et al. \(2013\)](#), [Li et al. \(2019\)](#), [Mei et al. \(2018\)](#), [Phan et al. \(2020\)](#), [Su et al., \(2019\)](#), [Tiwari et al. \(2019\)](#), [Zhenghui and Junhao \(2019\)](#).

<sup>3</sup> Data are extracted on March 10, 2020 from [http://www.policyuncertainty.com/us\\_daily.html](http://www.policyuncertainty.com/us_daily.html).

reported 19,461 new infection cases which created a panic on the US financial markets and the transactions were closed during that day. This number is intentionally omitted from Figure 1(b) which underpin two waves of COVID-19 new infections, the first recorded in China, and the second mirroring the conditions outside China, where countries as Italy, Iran and South Korea are severely affected. The death ratio climbed to 3.5% on March 10, 2020. Figure 1(c) shows that the number of infected people overpasses the psychological barrier of 100,000, whereas the number of affected countries increases exponentially since February 24, 2020. Finally, Figure 1(d) highlights opposite movements of oil prices and death ratio associated with COVID-19 starting with the second part of February.

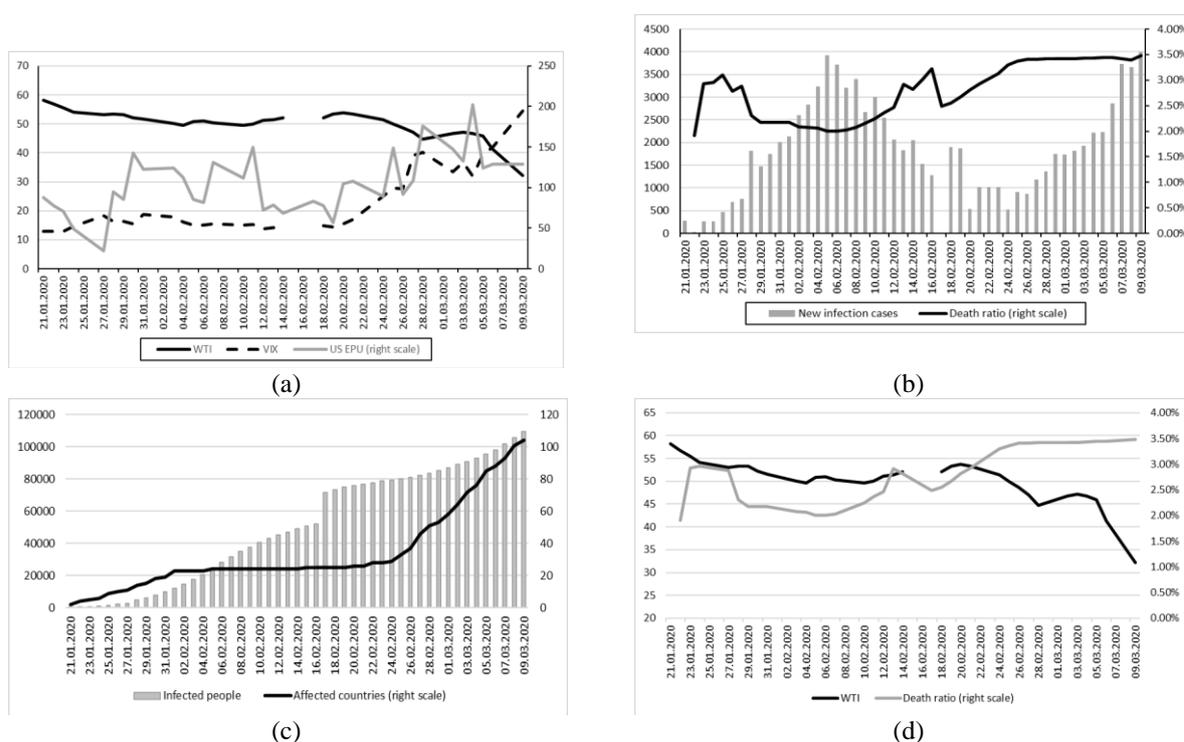


Fig. 1. Oil prices, financial volatility, economic uncertainty and coronavirus dynamics

Sources: WHO situation reports, Chicago Board Options Exchange (CBOE), Baker et al. (2016)- updates

The Phillips-Perron unit root test shows that our series are either I(1) or I(0) (Table 1).

Table 1. Unit root test

	Oil-WTI	COVID-19	VIX	EPU
Level	0.120	-6.228***	2.172	-4.785***
First difference	-3.561**	-39.39***	-5.254***	-27.61***

Notes: (i) \*\*\*, \*\* and \* means significance at 1%, 5% and 10%; (ii) the optimal lag selection is based on AIC information criterion; (iii) COVID-19 refers here to the new cases reported at global level.

## 2.2. Methodology

In order to estimate the relationship between oil prices, COVID-19, VIX and EPU, we use the ARDL model proposed by [Pesaran et al. \(2001\)](#), who is compatible with both I(0) and I(1) series.<sup>4</sup>

The [Pesaran et al.'s \(2001\)](#) framework uses a linear transformation to integrate short-run adjustments into the long-run equilibrium, resorting to an Error Correction Model (ERM), as follows:

$$\begin{aligned} \Delta Oil_t = & c + \delta_{oil} Oil_{t-1} + \delta_{COVID-19} COVID-19_t + \delta_{VIX} VIX_{t-1} + \delta_{EPU} EPU_{t-2} + \sum_{i=1}^p \alpha_i \Delta Oil_{t-i} + \\ & + \sum_{i=1}^p \beta_i \Delta COVID-19_{t-i} + \sum_{i=0}^p \gamma_i \Delta VIX_{t-i} + \sum_{i=1}^p \gamma_i \Delta EPU_{t-i} + \theta ECT_{t-i} + \varepsilon_t \end{aligned} \quad (1)$$

where: (i)  $c$  and  $\varepsilon$  is the intercept and the error term respectively, (ii) short-run terms are denoted by  $\Delta$ , whereas the long-run terms are indicated with  $\delta$ -terms, (iii)  $p$  is the maximum number of lags (four in our case), (iv) the error correction term is denoted by ECT ( $\theta$  should be negative and significant in order to validate the long-run relationship).

The existence of a long-run relationship is validated using a F-statistic, where the null hypothesis of no cointegration is  $\delta_{oil} = \delta_{COVID-19} = \delta_{VIX} = \delta_{EPU} = 0$ .<sup>5</sup>

## 3. Empirical results

In the first step we check for the existence of a long-run relationship, applying the bound tests which assume a lower bound for I(0) series, and an upper bound for I(1) series, the critical values being derived from [Narayan \(2005\)](#). The F-statistic indicates a cointegrating relationship if the values are higher than the critical value of the upper bound. For the first two models referring to the new infection cases daily reported at global level and in China, we notice the existence of a long-run relationship between oil prices, coronavirus numbers, financial volatility and the US economic uncertainty (Table 2). However, the situation outside China can only have a short-run effect on crude oil prices.

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<sup>4</sup> The WHO reports released at date “ $t$ ” contains information and figures about COVID-19 reported by countries at the end of day “ $t-1$ ”. The press usually announces new numbers, significantly higher at date “ $t$ ”, which are recorded by WHO reports at date “ $t+1$ ”. Therefore, in our estimation we consider the impact of COVID-19 data reported at “ $t+1$ ” on the oil price recorded at date “ $t$ ”. At the same time, the EPU index is daily updated, and might impact the oil prices with a small delay. Therefore,  $EPU_{t-1}$  is considered in our general equation.

<sup>5</sup> The optimal number of lags is chosen based on the Akaike Information Criteria (AIC) automatic selection criterion.

Table 2. Bounds test results (WTI)

Model specification	F-statistic	Critical values		Conclusion
		Lower bound (I(0))	Upper bound (I(1))	
COVID-19 Total	14.36	2.79	3.67	cointegration
COVID-19 China	4.004	2.79	3.67	cointegration
COVID-19 Outside China	1.607	2.79	3.67	no cointegration

Notes: (i) Critical values at 5% significance level.

In the second step we estimate the three ARDL models (Table 3). For Model 1, covering the overall situation, a negative long-run connection between oil prices, coronavirus numbers, financial volatility and US economic policy uncertainty is noticed. The effect is stronger for the VIX, whereas COVID-19 has only a marginal effect on WTI. A similar situation is recorded in the short run, except for the influence of US EPU, which becomes insignificant.

Table 3. Estimation of the ARDL specification (WTI)

	Model 1: COVID-19 Total		Model 2: COVID-19 China		Model 3: COVID-19 Outside China	
<b>Long-run equation</b>						
COVID-19 <sub>t+1</sub>	-0.001***	[0.000]	-0.001**	[0.000]		
VIX <sub>t</sub>	-0.282***	[0.014]	-0.200***	[0.044]		
EPU <sub>t-1</sub>	-0.009**	[0.004]	-0.060***	[0.011]		
c	5.987***	[0.310]	6.163***	[0.938]		
<b>Short-run equation</b>						
ΔOil <sub>t-1</sub>	0.536***	[0.108]			0.770**	[0.241]
ΔOil <sub>t-2</sub>	0.243*	[0.121]			-0.289	[0.151]
ΔOil <sub>t-3</sub>					0.498**	[0.142]
ΔCOVID-19 <sub>t+1</sub>	-0.000**	[0.000]	-0.000	[0.000]	-0.005***	[0.000]
ΔCOVID-19 <sub>t</sub>	-0.001***	[0.000]			0.000	[0.001]
ΔCOVID-19 <sub>t-1</sub>	0.001***	[0.000]			-0.003*	[0.001]
ΔCOVID-19 <sub>t-2</sub>	0.000**	[0.000]			-0.006**	[0.001]
ΔVIX <sub>t</sub>	-0.154***	[0.026]	-0.142***	[0.031]	-0.087**	[0.031]
ΔVIX <sub>t-1</sub>			-0.217***	[0.031]	-0.148**	[0.043]
ΔVIX <sub>t-2</sub>					0.120*	[0.056]
ΔEPU <sub>t-1</sub>	0.002	[0.002]	-0.017***	[0.004]	-0.022***	[0.004]
ΔEPU <sub>t-2</sub>			0.017**	[0.006]	0.016	[0.008]
ΔEPU <sub>t-3</sub>			0.021***	[0.005]	0.024	[0.007]
ΔEPU <sub>t-4</sub>			0.016***	[0.004]	0.014	[0.004]
ECT <sub>t-1</sub>	-1.453***	[0.151]	-0.821***	[0.161]	-0.721**	[0.189]
<b>Tests</b>						
Serial correlation	NO		NO		NO	
ARCH effects	NO		NO		NO	
Stability	YES		YES		YES	

Notes: (i) \*\*\*, \*\* and \* means significance at 1%, 5% and 10%; (ii) standard deviations are in square brackets; (iii) Breusch-Godfrey LM test for serial correlation is used; (iv) ARCH effects for conditional heteroscedasticity (with 4 lags); (v) Ramsey and CUSUM tests are used to check the stability.

Model 2 evidences similar findings in the long run, which is not surprising given that in January – February 2020, the new cases of infection are dominated by those reported in China. For Model 3, there is no significant long-run relationship between our variables. However, in the short run, it appears that the effect of COVID-19 on oil prices is more important as compared to those reported in China. These results can be easily explained by the reaction of commodity and financial markets to the COVID-19 spread in Europe and the US.

The robustness results considering the BRENT prices confirm our initial findings (Appendix). However, the existence of a cointegration relationship is documented at 10% significance level only (Table A1). The long-run relationship is significant in this case for Model 3, also. Nevertheless, Table A2 shows that in the long run, COVID-19 has a rather reduced negative impact on crude oil price, whereas the effect is not significant for the cases reported outside China. The short-run equation reveals mixed findings.

#### **4. Discussion and policy implications**

Even if the direct negative impact of COVID-19 new infection cases on crude oil prices is rather limited, the indirect effect manifested through the volatility of financial markets cannot be neglected. Given the speed of propagation of this pandemic virus, if the world governments are not proactive, promptly implementing the required measures to isolate the suspected cases of COVID-19, the entire economy will be paralyzed in few weeks.

The virus has already generated shock waves on financial markets, but also on commodity prices, including oil. Oil prices recorded the hardest cut after 1991 which help, for the moment, the economy of oil-importing countries severely affected by the coronavirus crisis. However, the crash of oil prices clearly shows that an economic downturn cannot be avoided. The amplitude of the economic contraction will be correlated with the coronavirus persistence. Although China seems to gain the fight against COVID-19, the virus exponentially propagates in Europe and the US. Consequently, a strong coordinated worldwide reaction is required, including economic measures to prevent a severe economic downturn. Central banks have already started to cut the interest rates, but this measure should be followed by appropriated fiscal facilities.

#### **5. Conclusion**

In the context of a severe oil price crash, the purpose of the paper was to see how the COVID-19 numbers, in terms of daily announcements of new infection cases, influenced the international oil prices. Our ARDL estimation documented a negative and significant impact of the coronavirus crisis, but relatively small as compared to the effect of financial volatility and economic policy uncertainty on oil prices.

Even if the figures reported outside China seem to have no significant effect on oil prices in the long run, the exponential increase of new infection cases risks to block the world economy and to freeze the oil prices at a low level for a long period.

### **Acknowledgements**

This work was supported by a Grant of the Romanian National Authority for Scientific Research and Innovation, CNCS–UEFISCDI, Project Number PN-III-P1-1.1-TE-2019-0436.

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## Appendix – Robustness results

Table A1. Bounds test results (BRENT)

Model specification	F-statistic	Critical values		Conclusion
		Lower bound (I(0))	Upper bound (I(1))	
COVID-19 Total	3.245	2.79	3.67	cointegration at 10% significance
COVID-19 China	2.853	2.79	3.67	inconclusive
COVID-19 Outside China	3.149	2.79	3.67	cointegration at 10% significance

Notes: (i) Critical values at 5% significance level.

Table A2. Estimation of the ARDL specification (BRENT)

	Model 1: COVID-19 Total		Model 2: COVID-19 China		Model 3: COVID-19 Outside China	
<b>Long-run equation</b>						
COVID-19 <sub>t+1</sub>	-0.001***	[0.000]	-0.001*	[0.001]	0.003	[0.001]
VIX <sub>t</sub>	-0.213***	[0.044]	-0.391*	[0.242]	-0.495**	[0.181]
EPU <sub>t-1</sub>	-0.041**	[0.014]	-0.056	[0.081]	-0.120***	[0.018]
c	6.663***	[0.705]	7.124***	[0.933]	7.546***	[0.363]
<b>Short-run equation</b>						
ΔOil <sub>t-1</sub>	0.560**	[0.239]	-0.351	[0.226]	0.094**	[0.571]
ΔOil <sub>t-2</sub>			-0.407	[0.269]		
ΔOil <sub>t-3</sub>			-0.183	[0.238]		
ΔCOVID-19 <sub>t+1</sub>	-0.000	[0.000]	-0.000	[0.000]	-0.000	[0.000]
ΔCOVID-19 <sub>t</sub>	-0.001***	[0.000]	-0.000	[0.000]	0.000**	[0.001]
ΔCOVID-19 <sub>t-1</sub>	0.001	[0.000]	-0.000	[0.000]	0.008***	[0.001]
ΔCOVID-19 <sub>t-2</sub>	0.000**	[0.000]			-0.003	[0.002]
ΔVIX <sub>t</sub>	-0.076	[0.057]	-0.173*	[0.079]	-0.060	[0.053]
ΔVIX <sub>t-1</sub>	-0.113	[0.066]	-0.268**	[0.091]	-0.068	[0.067]
ΔVIX <sub>t-2</sub>	0.088	[0.076]	-0.022	[0.114]	0.160*	[0.075]
ΔVIX <sub>t-3</sub>	-0.158*	[0.087]	0.015	[0.101]		
ΔEPU <sub>t-1</sub>	-0.015**	[0.006]	-0.014	[0.009]	-0.038***	[0.007]
ΔEPU <sub>t-2</sub>	0.028*	[0.013]	-0.012	[0.011]	0.041**	[0.014]
ΔEPU <sub>t-3</sub>	0.030***	[0.008]			0.055***	[0.013]
ΔEPU <sub>t-4</sub>					0.038***	[0.009]
ECT <sub>t-1</sub>	-1.443***	[0.292]	-0.401**	[0.158]	-0.876***	[0.180]
<b>Tests</b>						
Serial correlation	NO		NO		NO	
ARCH effects	NO		NO		NO	
Stability	YES		YES		YES	

Notes: (i) \*\*\*, \*\* and \* means significance at 1%, 5% and 10%; (ii) standard deviations are in square brackets; (iii) Breusch-Godfrey LM test for serial correlation is used; (iv) ARCH effects for conditional heteroscedasticity (with 4 lags); (v) Ramsey and CUSUM tests are used to check the stability.