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¹ A multivariate statistical strategy to adjust musculoskeletal models

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4 Abstract

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In musculoskeletal modelling, adjusting model parameters is challenging. This paper 5 proposes a multivariate statistical methodology to adjust muscle force-generating parame-6 ters optimally. Dynamic residuals are minimized as muscle force-generating parameters are 7 varied (maximal isometric force, optimal fiber length, tendon slack length and pennation an-8 gle). First, a sensitivity and a Pareto analyses are carried out in order to sort out and screen 9 the set of parameters having the greatest influence regarding the dynamic residuals. These 10 parameters are then used to create a response surface following a Design of Experiments 11 (DoE) approach. Finally, this surface is used to determine the optimum levels of the design 12 variables (muscle force-generating parameters). The proposed methodology is illustrated 13 by the adjustment of a three-dimensional musculoskeletal model of a sheep forelimb. After 14 adjustment, the reserve actuator values of the elbow and wrist joints were reduced, on aver-15 age, by 18%, and 16%, respectively. These results demonstrate that the use of multivariate 16 statistical strategies is an effective way to adjust model parameters optimally while reducing 17 dynamic inconsistencies. This study constitutes a step towards a more robust methodology 18 in musculoskeletal modelling, focusing on muscular parameter tuning. 19

- 20 Keywords: Musculoskeletal modelling, Force-generating parameters, Optimization,
- ²¹ Multivariate, Response surface

22 1. Introduction

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Musculoskeletal modelling allows estimating physical quantities that can be difficult to 23 measure (for instance muscle forces and joint reaction forces) or even studying what-if scenar-24 ios. A major challenge in this field is to adjust model parameters reliably, namely geometric 25 parameters, segment inertial parameters and muscle parameters. Geometric parameters can 26 be determined by scaling procedures (from generic models [23]), optimization-based pro-27 cedures (contact surfaces best-fit [20, 2]) or by functional movement analysis [31, 19, 32]. 28 There exist several methods in the literature either to experimentally estimate [1, 9, 18] or 29 to scale inertial parameters from anatomical datasets [23, 30]. Muscle parameters are typ-30 ically scaled from generic models [12, 34, 16] or estimated by functional approaches. Func-31 tional approaches require isometric trials using isokinetic ergometers or experimental ones 32 [27, 15, 13], which can be expensive, prone to modeling and experimental errors and difficult 33 to apply when dealing with animals. Modenese et al. [25] proposed a numerical method 34 for estimating the optimal fiber length and tendons slack length of the muscle-tendon units 35 by optimizing the operating range of the muscles simultaneously. In that approach, muscle 36 contraction dynamics are not required to be solved, making the method fast to compute. 37 However, in this approach, each muscle is solved independently, hindering the possibility to 38 solve the whole optimization problem as a multivariate one (pennation angle, tendon slack 39 length, optimal fiber length, maximal isometric force). 40

The effect of the uncertainty of musculoskeletal parameters has been evaluated by means of uncertainty propagation analysis [28, 29] and sensitivity analysis [4, 24, 11, 5]. These studies revel a high influence of the force-generating parameters (FGPs) in muscle force prediction (MFP). However, these parameters (pennation angle, tendon slack length, optimal fiber length, maximal isometric force) are difficult to measure or estimate [3, 25]; therefore, the use of numerical methods could be further explored as an option to adjust them.

⁴⁷ Using conventional optimization methods to adjust musculoskeletal models may become
⁴⁸ extremely time consuming due to the large number of optimization variables together with

the complexity of the MFP process (nonlinearities). Additionally, conventional optimization 49 methods typically vary one parameter at a time, thus the interdependence of variables is not 50 considered. Multivariate statistical approaches, such as Design of Experiments (DoE), offer 51 means for understanding the contribution of individual factors of a process to determine 52 the optimal operating conditions at low computational cost. This methodology consists in 53 fitting a multidimensional surface to the response variables through sequential variations of 54 the design variables over a specific region of interest. The parameters defining the response 55 surface can be determined by a least-squares regression. This surface can then be used 56 to determine the optimum levels of the design variables regarding the response variables. 57 DoE and response surface methodologies have been used in biomechanics [22, 8] but, to the 58 knowledge of the authors, it has not been explored as a possibility to adjust musculoskeletal 59 models. 60

⁶¹ Consequently, the focus of this paper is on the adjustment of the FGPs of the muscles
⁶² using a multivariate statistical methodology based on the DoE approach. Two algorithms are
⁶³ proposed in this work: one for screening the most influencing parameters, and a second aimed
⁶⁴ at creating the response surface and determining the optimum levels of the design variables.
⁶⁵ In the next section, the methodology to adjust FPGs is detailed. This methodology is
⁶⁶ applied to a musculoskeletal model of a sheep forelimb. The results are presented in Section
⁶⁷ 3 and discussed in Section 4.

68 2. Methods

The methodology proposed in this paper to adjust FGPs of the muscles is depicted in Figure 1. The methodology is based on three steps. The first step consists in screening the parameters by means of sensitivity analysis to identify those having the greatest impact on the performance of a given MFP. The performance of a simulation, as explained below, is defined in function of the dynamic residuals (reserve actuators). The second step consists Figure 1: Flow chart of the main steps of the proposed methodology to adjust musculoskeletal models. The methodology is based on three steps. The parameter screening step consists in computing the sensibility index S_k of the parameter P_k regarding the metric $\epsilon_{res}(t, \mathbf{P})$ (dynamic residuals). The response surface regression consist in fitting a multidimensional surface to the response variables by sequential variations of the influencing parameters. The model optimization step consists in finding the parameter values that minimizes the response variable (dynamic residuals).

⁷⁴ in constructing a response surface by means of the DoE approach. This surface represents ⁷⁵ the response of the simulation in terms of dynamic residuals to the variation of the model ⁷⁶ FGPs. The third step consists on the FGPs adjustment, looking for the minimum value of ⁷⁷ the response surface. This corresponds to find the values of the FGPs that minimize the ⁷⁸ dynamic residuals. In the following sections, the performance metric chosen to evaluate a ⁷⁹ MFP is presented, and then the three steps of the adjustment methodology are detailed.

80 2.1. MFP performance metric

For a given MFP process, let $\tau_j(t)$ be the net joint torque of the *j*-th joint (with $j = 1 : N_j joints$). Due to dynamic inconsistencies, the model would not be able to follow the desired kinematics with muscle actuators alone at a given time instant *t*. In order to allow the simulation to run, it is a common practice (which is the case of OpenSim, for example) to include additional actuators at each joint of the model R_j to reach the net joint torque. These are non-physical actuators that absorb dynamics residuals of the simulation. These actuators are added to each joint to reach the joint net torque:

$$\tau_j(t) = R_j + \sum_{i=1}^{N_{-muscles}} F_{0,i} \ r_{ij}(t) \ a_i(t) \ f(l_i(t)) \ g(v_i(t)) \ \cos(\phi_i(t)) \ \forall j \in [1; N_{-joints}]$$
(1)

where $F_{0,i}$ is the maximal isometric force of the *i*-th muscle (which is proportional to the physiological cross sectional area), r_{ij} is the moment arm of the *i*-th muscle about the *j*-th joint, $a_i(t)$ is the activation level (0 to 100%) of the *i*-th muscle at the time instant t, $f(l_i(t))$ is the force-length related coefficient, $g(v_i(t))$ is the force-velocity related coefficient and ϕ_i ⁹² is the pennation angle.

For each time instant of the simulation, these residuals R_j should be null if the MFP problem were perfectly solved. In consequence, it is beneficial to keep these values R_j at minimum levels. For each time instant t, the sum of the squares of reserve actuators of all joints of the model $\epsilon_{res}(t, \mathbf{P})$ has been chosen as the metric to evaluate the performance of an MFP:

$$\epsilon_{res}(t, \boldsymbol{P}) = \sum_{j=1}^{N_{-joints}} R_j^2 \tag{2}$$

where P is the vector containing all the FGPs of the model. This metric is used in this work to:

1. define the MFP sensitivity regarding variations in FGPs of the muscles,

¹⁰¹ 2. define the cost function used for the adjustment of the model.

102 2.2. Parameter screening

As it would be very time-consuming to adjust the complete set of FGPs of a musculoskeletal model, we propose first, by means of a sensitivity analysis, to screen the parameters having the greatest influence on the MFP performance. This process consists on quantifying the effects of varying each parameter value on the simulation performance. This performance is defined by means of the metric $\epsilon_{res}(t, \mathbf{P})$, as explained in the previous section. This process is summarized in Algorithm 1. Algorithm 1 Parameter screening

Require: Musculoskeletal model, kinematic data, external forces **Ensure:** Parameter sensibility 1: for each muscle M_i (with $i = 1 : N_{-}muscles$) of the model do 2:for each parameter P_k (with k = 1 : d) of M_i do for each increment step do 3: Increment the value of P_k 4: Estimate muscle forces for time $t = t_{start} : t_{end}$ 5:Get joint reserve values for time $t = t_{start} : t_{end}$ 6: for $(t = t_{start}; t < t_{end}; t = t + t_{step})$ do 7: Compute the metric value $\epsilon_{res}(t, \mathbf{P})$ (Eq. 2) 8: end for 9: end for 10: for $(t = t_{start}; t < t_{end}; t = t + t_{step})$ do Compute the numerical derivative $\frac{\partial \epsilon_{res}(t, \mathbf{P})}{\partial P_k}$ (Eq. 3) 11:12:end for 13:Compute sensitivity index S_k (Eq. 4) 14:Found the 20% of parameters that produce, at least, 80% of effects 15:end for 16:17: end for

The number of sensitivities to calculate is equal to the number of parameters to consider 109 per muscle times the number of muscles in the model. Each sensitivity calculation requires 110 estimating as many muscle force solutions as the number of points considered for the numer-111 ical differentiation. From a numerical point of view, using four points allows to smooth local 112 variations, improving the general estimate of the derivative value. In order to avoid this, we 113 used a four-point strategy (see Figure 2). The value of the k-th parameter of the model is 114 moved around its initial value P_k^* four times: $P_k^* - 2q_k$, $P_k^* - q_k$, $P_k^* + q_k$ and $P_k^* + 2q_k$. The 115 value of $2q_k$ corresponds to the variation coefficient and it has been chosen equal to 10%, as 116 in other works [35, 14]. 117

From the results of the four simulations for each time instant t, the sensitivity of each parameter P_k regarding the metric $\epsilon_{res}(t, \mathbf{P})$ can be estimated numerically using the fourpoints central difference approximation:

$$\frac{\partial \epsilon_{res}(t, \boldsymbol{P})}{\partial P_k} \approx -\frac{1}{12q_k} (8\epsilon_{res}(t, P_k^* - q_k) - 8\epsilon_{res}(t, P_k^* + q_k) + \epsilon_{res}(t, P_k^* + 2q_k) - \epsilon_{res}(t, P_k^* - 2q_k))$$
(3)

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Figure 2: Representation of the computation of $\partial \epsilon_{res}(t, \mathbf{P})/\partial P_k$. At the left-hand side it is represented the evolution of the metric value $\epsilon_{res}(t, \mathbf{P})$ in function of the time t and the parameter value P_k . At the right-hand side, the sensitivity of $\epsilon_{res}(t, \mathbf{P})$ regarding variations on the P_k around a given value P_k^* is approximated by a four-point numerical derivative. The variation coefficient is $2q_k$.

As Eq. 3 is formulated for a given instant t, a set of sensitivities is obtained when considering all time instants from t_{start} to t_{end} of the simulation. From these values, a global sensitivity index is defined for the k-th parameter as follows:

$$S_k = max\left(\frac{\partial \epsilon_{res}(t, \boldsymbol{P})}{\partial P_k}\right) - min\left(\frac{\partial \epsilon_{res}(t, \boldsymbol{P})}{\partial P_k}\right), \ \forall t \in [t_{start}; t_{end}]$$
(4)

In order to determine the set of FGPs having the greatest sensitivity, a Pareto heuristic (also known as the 80/20 rule) can be used. This heuristic, which is used in decision making, suggests that 20% of parameters produce 80% of the effects. In our case, these effects correspond to the cumulative sum of sensitivity factors. These FGPs are those that are going to be adjusted for reducing the dynamic inconsistencies of the computation, as explained in the next section.

¹³¹ 2.3. Response surface regression and model optimization

The set of FGPs identified as significant after screening are those chosen for optimization. The cumulative sum of squared reserve actuators was chosen as the cost function to be minimized:

Given
$$\boldsymbol{P} = [P_1, P_2, ..., P_d]$$

minimize $C_{res}(\boldsymbol{P}) = \sum_{t=t_{start}}^{t_{end}} \epsilon_{res}(t, \boldsymbol{P})$ (5)
with $t = t + t_{step}$

135 where t_{step} is the time frame of the experimental data.

The response surface methodology was used to perform this optimization. This method-136 ology consists in generating an empirical mathematical model that relates some design vari-137 ables with a response variable. A hypersurface is fitted to the response variable on a region 138 of interest by a systematic variation of the design variables. In our case, the design variables 139 (or factors) correspond to the model FGPs $\boldsymbol{P} = [P_1, P_2, ..., P_d]$ and the response variable to 140 the dynamic residuals C_{res} . This surface is used to find the design variable values (FGPs) 141 that minimize the response variable (dynamic residuals). The definition of the response 142 surface and the optimization process are summarized in Algorithm 2 and explained below. 143

Algorithm 2 Parameter adjustment via response surface regression

Require: Musculoskeletal model, kinematic data, **Require:** external forces, set of influencing FGPs

Ensure: Set of optimal FGP values

1: d = get the number of influencing FGPs (from §2.2)

- 2: Create the design matrix
- 3: for run = 1 : m do

4: Modify the P_k value $\forall k \in [1:d]$ according to the experiment design

- 5: Estimate muscle forces for time $t = t_{start} : t_{end}$
- 6: Get joint reserve values for time $t = t_{start} : t_{end}$
- 7: **for** $(t = t_{start}; t < t_{end}; t = t + t_{step})$ **do**
- 8: Compute the metric value $\epsilon_{res}(t, \mathbf{P})$ (Eq. 2)
- 9: end for

10: Compute the sum of the squares of reserves: $\mathbf{y}[run] = C_{res}(\mathbf{P})$ (Eq. 2.3)

- 11: Compute the surface response coefficient vector \mathbf{b} (Eq. 9)
- 12: Find $\mathbf{x}_{\min} = [x_0, \dots, x_d]$ that minimizes $y(x_0, \dots, x_d)$
- 13: Evaluate the predicted results by means of the $SCB(\hat{y})$ (Eqs. 13 and 14)
- 14: Validate the predicted results
- 15: end for

In the response surface methodology, a typically employed model is a second-order hypersurface $y(x_0, \ldots, x_d) \in \mathbb{R}^d$:

$$y(x_0, \dots, x_d) = \sum_{k=0}^d \sum_{j=k}^d x_k x_j b_{kj}$$
(6)

where d corresponds to the number of FGPs chosen for the experiment after screening, b_{kj} are the surface coefficients, $x_0 = 1$ and x_k is the normalized value of the parameter P_k (with k = 1 : d).

The number of terms in Eq. 6 is defined by both the order of the surface (2 in our case) and the dimension d of the space, and it is n = C(d + 2, 2) = (d + 2)!/(2!d!). Hence, in order to determine the n coefficients that define the surface, it is necessary to run at least nexperiments.

According to DoE methodology, more than n experiments are usually executed. For the execution of these experiments, the values of the design variables (factor levels) are systematically moved around an initial value. These initial values, which are assumed to be near to a locally optimal solution, define the central point of the experiment. The variation of the values of these parameters around the central point is defined by the design type.

For fitting a second-order model, a central composite design (CCD) type is a popular 158 choice. This type of design defines the same number of levels for the different factors. For 159 each factor, five levels are considered. Factor levels are normalized in such a way that their 160 inputs were coded as +1, $+1/\sqrt{d}$, 0, $-1/\sqrt{d}$ and -1. A CCD can be constructed using a full or 161 a fractional factorial design. Choosing a full or a fractional factorial design type depends on 162 the number of factors to consider. This choice defines the number of experiments to perform 163 and the fidelity (or the uncertainty) of the empirical model regarding the represented process 164 (see [7] for more details). For a full factorial design, which is recommended when $d \leq 4$, the 165 number of required runs is $m = 2^d + 2d + 1$. For d > 4, a half factorial design can be used 166 (number of runs: $m = 2^{(d-1)} + 2d + 1$). 167

According to the design type and the number of factors, the design matrix can be obtained. This matrix provides a program for the m experiments by defining the normalized values of the x_k factors (with k = 1 : d) for each run. This matrix can be obtained, for example, using the Matlab (Mathworks Inc.) function *ccdesign*. From the execution of the m experiments, a set of m values of the response variable are obtained:

$$\begin{bmatrix} y_1 \\ \vdots \\ y_m \end{bmatrix} = \begin{bmatrix} 1 & x_{11} & \dots & x_{1n} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{m1} & \dots & x_{mn} \end{bmatrix} \begin{bmatrix} b_0 \\ \vdots \\ b_n \end{bmatrix}$$
(7)

¹⁷³ In matrix form, Eq. 7 becomes:

$$\mathbf{y} = \mathbf{X}\mathbf{b} \tag{8}$$

where **y** is the response vector $(m \times 1)$, the matrix **X** $(m \times n)$ is called matrix of independent variables and it is obtained from the design matrix and the vector **b** $(n \times 1)$ contains the coefficients b_{kj} of the response surface. As **X** is in general a non-square matrix (m > n), one can compute an equivalent vector $\hat{\mathbf{b}}$ that best estimate of the surface response coefficients using the pseudo-inverse method as follows:

$$\hat{\mathbf{b}} = (\mathbf{X}^{\mathrm{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathrm{T}}\mathbf{y}$$
(9)

From the calculated surface, a local minimum $y_{min} = C_{res}^{pre}$ can be predicted, which is to search for a set of normalized parameter values $\mathbf{x_{min}} = [x_0, \dots, x_d]$ that minimizes the response value $y(x_0, \dots, x_d)$. This is a nonlinear optimization that can be soled by using, for example, the generalized reduced gradient method [21].

In order to asses the quality (or inaccuracy) of the fit of the surface regarding the experimental values can be assessed by means of the root mean square (RMS) of the residues (Eq. 10). These residues are calculated as the difference between the experimental results (y) and the corresponding points on the response surface $(X\hat{b})$:

$$E = \text{RMS}(\mathbf{y} - \mathbf{X}\hat{\mathbf{b}}) \tag{10}$$

The *E* value can also be used to calculate the statistical confidence boundary (SCB) (or error bar) of a predicted value \hat{y} . This is based on classical uncertainty propagation method as explained as follows (for more details refer to [6]).

In order to calculate the SCB value at a given point, it is required to start calculating the covariance matrix of $\hat{\mathbf{b}}$ by the following expression:

$$\mathbf{cov}(\hat{\mathbf{b}}) = (\mathbf{X}^{\mathrm{T}}\mathbf{X})^{-1}E^2 \tag{11}$$

The deviation of the response surface at a given point \hat{y} can be evaluated by calculating the mean square error (MSE):

$$MSE(\hat{y}) = \mathbf{J} \operatorname{cov}(\hat{\mathbf{b}}) \mathbf{J}^{\mathrm{T}}$$
(12)

where **J** is the Jacobian of the function $y = f(x_0, ..., x_d)$ (Eq. 6) with respect to the b_{kj} coefficients. In other words, **J** is the vector of the partial derivatives $\frac{\partial y}{\partial b_{kj}}$. Using equation 12, the SCB of the surface at a given point \hat{y} can be calculated:

$$SCB(\hat{y}) = \hat{y} \pm k_n \sqrt{MSE(\hat{y})}$$
 (13)

where k_n is calculated using a normal law distribution for a given two-sided confidence level. Even if the response variable C_{res} is based on a squared value (Eqs. 2.3 and 2), the lower bound of the SBC of a given predicted value \hat{y} could be negative. In this case, the SBC value must be corrected using a folded normal distribution, as follows:

$$SCB(\hat{y}) = 0 + k_r \sqrt{MSE(\hat{y})}$$
(14)

where k_r is calculated using the folded normal distribution for a given one-sided confidence level.

203 2.4. Case study: adjustment of a sheep forelimb model

In order to illustrate the proposed approach, a sheep forelimb musculoskeletal model was used. The musculoskeletal model was composed by 4 segments and 16 hill-type muscletendon units (see Figure 3). The external geometry of the bones was obtained by 3D optical scanning. The extensor and flexor groups of the digital joints were treated each one as a whole unit as no degree of freedom at these joints were considered. Elbow and carpal joints were modeled as hinge joints and shoulder joint as a ball-and-socket joint. Inertial parameters were estimated using a CAD software. A geometrical model of the segments was created by considering all soft tissue (see Figure 4). The model and the gait trial data used in this case study are available at https://simtk.org/projects/sheepforelimb. The considerations made for carrying out this study were:

214	• Bones were considered as rigid bodies,
215	• Segments' inertial properties were considered time-independent,
216	• Biological joints were modelled as frictionless joints,

• The initial values of the FGPs were considered near to a locally optimal solution.

Figure 3: Musculoskeletal model of sheep forelimb used as case study. The musculoskeletal model was composed by 4 segments and 16 hill-type muscle-tendon units (represented as red lines) Anconaeus (ACN), Biceps brachii (BCB), Brachialis (BRC), Caput mediale tricipitis brachii (MET), Caput laterale tricipitis brachii (LAT), Caput longum tricipitis brachii (LOT), Coracobrachialis (CCB), Deltoideus (DET), Extensor group (EXT), Flexor group (FLE), Infraspinatus (ISP), Subscapularis (SUB), Supraspinatus (SUP), Tensor fascie antebranchii (TFA), Teres major (TMA), Teres minor (TMI).

Figure 4: Computation of the inertial parameters (inertia matrix, mass and center of gravity) of the model's segments using a CAD software. For each segment, soft tissues were considered.

The open source tool OpenSim [10, 33] was used to perform MFP of the gait cycle. For 218 these computations, the static optimization algorithm was used. The algorithms proposed in 219 this paper were implemented in Jython and the computations were automatically performed 220 through the OpenSim API on a standard laptop (64-bit i7 processor @ 2.7GHz, 16 GB RAM, 221 running Microsoft Windows 7). For the screening process, four FGPs per muscle (those 222 assumed to be specific to each actuator) were considered, namely maximal isometric force, 223 optimal fiber length, tendon slack length and pennation angle. For the numerical derivative 224 computations, a variation coefficient of $2q_k = 0.1P_k^*$, $\forall k$ was chosen. This is a commonly 225 used value in biomechanics [35, 14]. According to Algorithm 1, that implies running $16 \times 4 \times 4$ 226 computations to calculate each sensitivity index. Muscle forces were computed before and 227

²²⁸ after FGPs adjustment. Reserve actuator values were compared with the net joint torques ²²⁹ for each time instant from the results of the two computations.

230 3. Results

231 3.1. Initial simulation

When computing muscle forces with the initial set of FGP values, a cumulative sum of 232 squared reserve actuators of $C_{res}^{ini} = 21.385e-3 \ (BW m)^2$ was obtained. This value was nor-233 malized according to the body weight (BW) of the specimen. Figure 5 shows the behaviour 234 of these reserves during the whole computation at the shoulder, elbow and wrist regarding 235 the net torque of each joint. At the shoulder, the peak reserve actuator was 9% of the peak 236 joint torque. At the elbow, the peak reserve actuator was 51% of the peak joint torque. 237 And at the wrist, the peak reserve actuator was 100% of the peak joint torque. Following 238 modelling good practices [17], these values must be less than 10%, which indicates that the 239 model must be adjusted. 240

(b) Elbow

(c) Wrist

Figure 5: Evolution of the joint net torque (continuous line) for the three joints during the simulation in comparison with the reserve actuators required to run the simulation with the initial model (dashed line) and with the adjusted model (dotted line): (a) for the shoulder joint, (b) for the elbow joint, (c) for the wrist joint.

²⁴¹ 3.2. Parameter screening

For screening purposes, a total of 256 MFP were executed, which took approximately 40 minutes. These results are summarized in Figure 6 by means of a Pareto chart. This chart represents the sensitivity factors sorted by magnitude. The curve over the bars represents the percentage of the cumulative sum of sensitivity factors. By means of this chart, the most influencing model parameters can be identified. We chose a set of 9 FGPs, that corresponds to 14% of the whole set. This set generates 97% of the sum of the sensitivity factors. As it can be noticed, among the whole set of muscle-tendon units of the model, three have a high influence on the dynamic consistency of the simulation: the extensor group of the digital joints (EXT), the biceps brachii (BCB) and the brachialis (BRC). Among the four FGPs considered, muscle pennation angles had less influence.

252 3.3. Surface regression

A response surface was calculated in the neighborhood of the initial values of these FGPs. The computation of this response surface required simulating 275 $(2^{(9-1)+2\times9+1})$ gait trials, which took approximately 48 minutes. The obtained surface coefficients are presented in Table 1. The convexity of a second-order surface depends on the sign of the second-order coefficients. As it can seen, all the second order coefficients are positive, which indicates that the obtained surface is convex.

Figure 6: Pareto chart of the sensitivity analysis. FGPs are represented in the horizontal axis sorted its own sensitivity index (vertical left-hand axis). The curve over the bars represents the percentage of the cumulative sum of sensitivity factors (vertical right-hand axis).

k	b_{0k}	b_{1k}	b_{2k}	b_{3k}	b_{4k}	b_{5k}	b_{6k}	b_{7k}	b_{8k}	b_{9k}
0	5.24									
1	-7.72	3.72								
2	-6.85	9.48	3.72							
3	-11.01	7.56	6.64	2.70						
4	-7.48	4.95	4.35	8.16	2.90					
5	-6.32	4.16	3.66	6.85	4.29	2.48				
6	-2.81	1.90	1.66	1.32	2.09	1.75	1.53			
7	-0.67	0.47	0.41	0.83	-0.58	-0.49	0.21	1.36		
8	-0.67	0.50	0.44	-0.74	0.59	0.49	-0.14	0.06	1.36	
9	-0.37	0.01	0.01	0.36	0.23	0.19	0.09	0.03	0.03	1.38

Table 1: Set of 55 b_{kj} coefficients of the response surface calculated from the results of the experiments by least-squares fitting.

259 3.4. Validation of the results

The obtained surface predicted an optimal design solution at $C_{res}^{pre} = 0$ (BW m)² with a statistical confidence boundary of $0 \leq \text{SCB}(C_{res}^{pre}) < 4.763\text{e}-3$ (BW m)². The confidence boundary was calculated considering a folded normal distribution (Eq. 14) and it is depicted in Figure 7.

Figure 7: Folded normal distribution used to calculate the corrected statistical confidence boundary of the predicted value C_{res}^{pre} for a confidence level of 95%.

The adjusted FGP values (shown in Table 2) were introduced in the model and the MFP 264 executed again. The obtained sum of reserve actuators was significantly reduced compared 265 to the computation performed using the initial set of FGP values. The new computed value, 266 which is included in the confidence boundary, was of $C_{res}^{cmp} = 0.120e-3 \text{ (BW m)}^2$. Reserve 267 actuator values of the simulation with the adjusted model are presented in Figure 5. After 268 adjustment, peak reserve values reached a maximum of 9%, 11% and 23% of the peak joint 269 torque of the shoulder, elbow and wrist, respectively. The average reduction of the reserve 270 value vs the net joint torque all along the simulation was of 1% for the shoulder, 18% for 271 the elbow and 16% for the wrist. 272

From both the initial model and the adjusted one, we computed the radiohumeral reaction force (Figure 8). The magnitude of this force in both axial and anterior-posterior directions shows differences before and after the FGPs adjustment. These differences are more significant especially during the stance phase.

(a)

(b)

Figure 8: Radiohumeral reaction forces predicted from the sheep forelimb musculoskeletal model before (continuous line) and after FGPs adjustment (dashed line). Force values were normalized according to the body weight (BW) of the specimen. (a) axial force and (b) anterior-posterior force

Variable	Unit	Initial values	Adjusted values
Opt. Fiber Len. EXT	m	0.0900	0.1116
TendonLen EXT	m	0.1700	0.1508
Opt. Fiber Len. BRC	m	0.0680	0.0843
Opt. Fiber Len. BCB	m	0.0680	0.0843
Tendon Len BCB	m	0.0910	0.1128
Tendon Len BRC	m	0.0340	0.0421
Max Force BCB	Ν	488.1200	509.1856
Max Force BRC	Ν	361.5700	375.1500
Max Force EXT	Ν	1193.1820	1201.3149

Table 2: Set of initial and adjusted values of the 9 FGPs selected after screening.

277 4. Discussion and conclusions

In response to the multiple challenges in estimating model parameters in musculoskeletal modelling, this paper proposes a multivariate statistical strategy to adjust FGPs. The methodology follows a DoE-based approach to generate a response surface, which is used to determine optimal levels of the FGPs. In contrast to conventional optimization methods, DoE-based methodologies offer a means to determine the optimal operating conditions of a process at low computational cost while considering the interdependence of the factors.

We found that the use of response surfaces in musculoskeletal dynamics computations 284 is an effective way to optimally adjust model FGPs, making the model more dynamically 285 consistent. In agreement with the findings of Modenese et al. [26], the results of our com-286 putation suggest that the reserve actuators have an impact on the predicted joint reaction 287 forces. Such an impact reached, in our case study, magnitudes of up to 1.7 BW. As re-288 serve actuators are applied directly over the joint (as a pure torque) without a moment 289 arm, they do not generate reaction forces at the joints. In consequence, if a joint reaction 290 analysis is performed with high reserve actuator values the calculated reaction forces are 291 underestimated. 292

Depending on the cost function used to generate the response surface, this methodology can be used to adjust geometrical, inertial or muscle parameters. The focus of this paper is ²⁹⁵ on the adjustment of FGPs to reduce joint reserve actuators. However, its application for a ²⁹⁶ geometrical or a inertial adjustment is straightforward.

It should be stressed out that the proposed method does not ensure the robustness of the model, but it does help in the adjustment of the parameters during the model development process. Furthermore, a model adjustment carried out following this approach is valid for the used experimental data and it may not be for another dataset.

This work has several limitations. First, as in the case of most of optimization techniques, 301 the response surface methodology is very sensitive to the initial values of the optimization 302 variables. Thus, the method proposed in this study should be employed using a realistic 303 set of initial FGP values. During the optimization, keeping parameter values in physical 304 levels depends on the chosen variation coefficient. Having small variation levels prevents 305 the optimization process to allocate a non-physical value to a given parameter. Second, 306 the proposed methodology could become impractical when the number of parameters to 307 optimize is significant (greater than 15), hence the importance of the parameter screening 308 before the optimization. Future work is in progress to develop a DoE-based strategy that 309 uses several locomotion trials for the adjustment of the FGPs. 310

This study showed that the use of multivariate statistical strategy seems to be a good alternative to optimally adjust FGPs. This contribution constitutes a step towards a more robust methodology in musculoskeletal modelling, methodology that can be applied both to human and non-human studies.

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320 Conflict of interest statement

³²¹ The authors declare no conflict of interest.

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