HAL open science

Point Processes of Non stationary Sequences Generated by Sequential and Random Dynamical Systems

Ana Cristina Moreira Freitas, Jorge Milhazes Freitas, Mário Magalhães, Sandro Vaienti

▶ **To cite this version:**

Ana Cristina Moreira Freitas, Jorge Milhazes Freitas, Mário Magalhães, Sandro Vaienti. Point Processes of Non stationary Sequences Generated by Sequential and Random Dynamical Systems. Journal of Statistical Physics, 2020, 181 (4), pp.1365-1409. 10.1007/s10955-020-02630-z . hal-02450476

HAL Id: hal-02450476

https://hal.science/hal-02450476v1

Submitted on 23 Jan 2020

HAL is a multi-disciplinary open access archive for the deposit and dissemination of scientific research documents, whether they are published or not. The documents may come from teaching and research institutions in France or abroad, or from public or private research centers.

L'archive ouverte pluridisciplinaire **HAL**, est destinée au dépôt et à la diffusion de documents scientifiques de niveau recherche, publiés ou non, émanant des établissements d'enseignement et de recherche français ou étrangers, des laboratoires publics ou privés.

POINT PROCESSES OF NON STATIONARY SEQUENCES GENERATED BY SEQUENTIAL AND RANDOM DYNAMICAL SYSTEMS

ANA CRISTINA MOREIRA FREITAS, JORGE MILHAZES FREITAS, MÁRIO MAGALHÃES, AND SANDRO VAIENTI

Abstract. We give general sufficient conditions to prove the convergence of marked point processes that keep record of the occurrence of rare events and of their impact for non-autonomous dynamical systems. We apply the results to sequential dynamical systems associated to uniformly expanding maps and to random dynamical systems given by fibred Lasota Yorke maps.

1. Introduction

The study of dynamical systems, initiated with Poincaré, concerns the qualitative description of the orbits of a system. A discrete system is defined by three main ingredients: a phase space, whose elements describe the state of the system at any given moment; the time, which is usually modelled by $\mathbb{Z}$ or $\mathbb{N}$; and a law that rules the action of time in the system, which is usually given by a map that determines the transition from a given state to a successive one. The orbits are then defined by an initial state and the subsequent states determined by the evolution law that rules the system. In the case the map is invertible, the orbit also includes the predecessing states. The main goal is to understand the time evolution of the orbits and in particular their asymptotic behaviour.

The emergence of chaotic systems brought renewed interest and opened new lines of research. In fact, sensitivity to initial conditions, *i.e.*, the fact that two orbits become uncorrelated regardless of how close the initial states are, introduced an unpredictability which made the classical analytical tools ineffective to understand the global behaviour of such systems. The complexity of the orbital structure of chaotic systems brought special attention to the study of limiting laws of stochastic processes arising from such systems, since they borrow at least some probabilistic predictability to their erratic behaviour. These stochastic processes arise in a very natural way simply by evaluating a given observable function along the orbits of the system.

The first step in this research direction is usually the construction of invariant physical measures, which provide an asymptotic spatial distribution of the orbits in the phase space and endow the stochastic processes dynamically generated with stationarity. Ergodicity then gives strong laws of large numbers. The mixing properties of the system restore asymptotic independence and, in this way, allow to mimic independent and identically distributed (iid) processes and prove

2010 *Mathematics Subject Classification.* 37A50, 60G70, 60G57, 37B20.

MM was partially supported by FCT grant SFRH/BPD/89474/2012, which is supported by the program POPH/FSE. ACMF and JMF were partially supported by FCT projects FAPESP/19805/2014 and PTDC/MAT-CAL/3884/2014, PTDC/MAT-PUR/28177/2017, with national funds. All authors would like to thank the support of CMUP (UID/MAT/00144/2013), which is funded by FCT with national (MCTES) and European structural funds through the programs FEDER, under the partnership agreement PT2020. JMF would like to thank the University of Toulon for the appointment as "Visiting Professor" during the year 2018.

limiting laws for the mean, such as: central limit theorems, large deviation principles, invariance principles, among others. However, in many occasions the exact formula for the invariant measure is not available and one has to rely on reference measures with respect to which these processes are not stationary anymore. Loosening stationarity leads to non-autonomous dynamical systems for which the study of limit theorems is just at the beginning. We mention the recent works [AHN+15, HNTV17, NTV18] and references therein.

While the limiting laws mentioned so far pertain to the mean or average behaviour of the system, in the recent years, the study of the extremal behaviour, ie, the laws that rule the appearance of abnormal observations along the orbits of the system has suffered an unprecedented development ([LFF+16]). In fact, Extreme Value Theory (EVT) became a new powerful tool to investigate the statistical properties of dynamical systems. The first approach followed in the beginning was to look at the partial maxima of the sequence of random variables generated by the observations along the orbits. It turned out that for particular choices of the observable function, the asymptotic distribution of the maximum gave new and important insight about the recurrence properties of the systems. This intimate connection between the extremal properties of the systems and their recurrence properties was established in [Col01, FFT10, FFT11]. Based on an extremal analysis of the systems, one can compute the probability of the first visit to shrinking targets on the phase space and show that it is determined by the local behaviour of the invariant measure. Moreover, the quantities associated with statistics of extremes, such as the *tail index* or the *extremal index*, revealed to be very useful to provide relevant information about, for example, the local dimension of the invariant measure, the expansion rate at periodic points, the Lyapunov exponents or the compatibility between the dynamics and the fractal structure of the target sets ([LFTV12, FFT12, CFM+18, FFRS18]).

We would like to stress that these remarkable progress in the description and the understanding of recurrence and the extremal behaviour of dynamical systems are not mere applications of the classical EVT established in probability. The very nature of the processes arising from the dynamical systems, which are far from being iid, addresses new theoretical issues and enlarges the horizon of the theory itself.

The study of rare (extreme) events for dynamical systems was initially performed under stationarity. Very recently, in [FFV17, FFV18], the authors developed tools to obtain the limiting distribution for the partial maxima of non-stationary stochastic processes arising from sequential dynamical systems ([BB84, CR07]) and random transformations or randomly perturbed systems ([Kal86, Kif88]). In the case of random transformations, we also mention the papers [RSV14, Rou14, RT15], where limiting laws for the waiting time to hit/return to shrinking target sets in the phase space (which are related to the existence of limiting laws for the maximum) were obtained for random dynamical systems.

The main purpose of this paper is to enhance the study of rare events for non-autonomous systems and, therefore, in a non-stationary context, by considering the convergence of point processes instead of the more particular distributional limiting properties of the maximum or the hitting/return times statistics. Point processes have revealed as a powerful tool to study the extremal behaviour of stationary systems. The most simple point processes, the Rare Events Point Processes (REPP) keep track of the number of exceedances (abnormally high values) observed along the orbits of the system and allow to recover relevant information such as the expected time between the occurrence of extremal events, the intensity of clustering, the distribution of the higher order statistics such as the maximum. For stationary systems, they were studied in [FFT13]. We

will also consider more sophisticated Marked Point Processes of Rare Events (which are random measures), studied for autonomous systems in [FFMa18] and which not only keep track of the number of exceedances but also of their impact. In the presence of clustering of rare events, we will be particularly interested in Area Over Threshold (AOT) marked point processes, which sum all the excesses over a certain threshold within a cluster, and Peak Over Threshold (POT) marked point point processes, which consider the record impact of the highest exceedance by taking the maximum excess within a cluster. The first allows to study the effect of aggregate damage, while the second focuses on the sensitivity to very high impacts. The potential of interest of these results is quite transversal, but we mention particularly the possible applications to climate dynamics where the study of extreme events for dynamical systems have proved to be very useful in the analysis of meteorological data (see for example [SKF$^+$16, MCF17, MCB$^+$18, FACM$^+$19]).

The paper is structured as follows. In Section 2, we generalise the theory developed in [FFV17] in order to obtain the convergence of Marked Point Processes of Rare Events (MREPP). In particular, we introduce the notation, concepts and conditions that allow us to state a result that establishes the convergence of the MREPP to a compound Poisson process for non-stationary stochastic processes, under some amenable conditions designed for application to non-autonomous systems. We believe that formula (2.13) which gives the multiplicity distribution of the limiting compound Poisson process has an interest on its own. Section 3 is dedicated to the proof of the main convergence result stated in the previous section. In Section 5, we make a non-trivial application of our main convergence result to some sequential dynamical systems studied in [CR07], deriving exact formulas for the limiting multiplicity distribution. In Section 6, we establish a convergence limiting result of the MREPP in the random dynamical systems setting, where we consider fibred LasotaYorke maps which were introduced in the recent paper [DFGTV18].

2. The setting and statement of results

Let $X_0, X_1, \ldots$ be a stochastic process, where each r.v. $X_i : \mathcal{Y} \to \mathbb{R} \cup \{\pm\infty\}$ is defined on the measure space $(\mathcal{Y}, \mathcal{B}, \mathbb{P})$. We assume that $\mathcal{Y}$ is a sequence space with a natural product structure so that each possible realisation of the stochastic process corresponds to a unique element of $\mathcal{Y}$ and there exists a measurable map $\mathcal{T} : \mathcal{Y} \to \mathcal{Y}$, the time evolution map, which can be seen as the passage of one unit of time, so that

$$X_{i-1} \circ \mathcal{T} = X_i, \quad \text{for all } i \in \mathbb{N}.$$

The σ-algebra $\mathcal{B}$ can also be seen as a product σ-algebra adapted to the X_i's. For the purpose of this paper, $X_0, X_1, \ldots$ is possibly non-stationary. Stationarity would mean that $\mathbb{P}$ is $\mathcal{T}$-invariant. Note that $X_i = X_0 \circ \mathcal{T}_i$, for all $i \in \mathbb{N}_0$, where $\mathcal{T}_i$ denotes the i-fold composition of $\mathcal{T}$, with the convention that $\mathcal{T}_0$ denotes the identity map on $\mathcal{Y}$. In the applications below to sequential dynamical systems, we will have that $\mathcal{T}_i = T_i \circ \ldots \circ T_1$ will be the concatenation of i possibly different transformations $T_1, \ldots, T_i$, so that

$$X_n = \varphi \circ \mathcal{T}_n, \quad \text{for all } n \in \mathbb{N} \tag{2.1}$$

for some given observable $\varphi : \mathcal{Y} \to \mathbb{R} \cup \{\pm\infty\}$

Each random variable X_i has a marginal distribution function (d.f.) denoted by F_i, *i.e.*, $F_i(x) = \mathbb{P}(X_i \leq x)$. Note that the F_i, with $i \in \mathbb{N}_0$, may all be distinct from each other. For a d.f. F we let $\bar{F} = 1 - F$. We define $u_{F_i} = \sup\{x : F_i(x) < 1\}$ and let $F_i(u_{F_i}-) := \lim_{h \to 0^+} F_i(u_{F_i} - h) = 1$

for all i. We will consider the limiting law of

$$\mathbf{P}_{H,n} := \mathbb{P}(X_0 \le u_{n,0}, X_1 \le u_{n,1}, \dots, X_{Hn-1} \le u_{n,Hn-1})$$

as $n \to \infty$, where $\{u_{n,i}, i \le Hn-1, n \ge 1\}$ is considered a real-valued boundary, with $H \in \mathbb{N}$.

We assume throughout the paper that

$$\bar{F}_{n,\max}(H) := \max\{\bar{F}_i(u_{n,i}), i \le Hn-1\} \to 0, \quad \text{as } n \to \infty, \tag{2.2}$$

and, for some $\tau > 0$,

$$\sum_{i=0}^{h_n-1} \bar{F}_i(u_{n,i}) = \frac{h_n}{n}\tau + o(1), \tag{2.3}$$

for any unbounded increasing sequence of positive integers $h_n \le Hn$. In particular, we have

$$F^*_{H,n} := \sum_{i=0}^{Hn-1} \bar{F}_i(u_{n,i}) \to H\tau, \quad \text{as } n \to \infty. \tag{2.4}$$

The most simple point processes that we will consider here keep track of the exceedances of the high thresholds $u_{n,i}$ by counting the number of such exceedances on a rescaled time interval. These thresholds are chosen such that

$$F^*_{1,n} = \sum_{i=0}^{n-1} \bar{F}_i(u_{n,i}) \to \tau, \quad \text{as } n \to \infty, \tag{2.5}$$

so that the average number of exceedances among the first n observations is kept, approximately, at the constant frequency $\tau > 0$.

2.1. **Random measures and weak convergence.** We start by introducing the notions of *random measures* and, in particular, *point processes* and *marked point processes*. One could introduce these concepts on general locally compact topological spaces with countable basis, but we will restrict to the case of the positive real line $[0,\infty)$ equipped with its Borel σ-algebra $\mathcal{B}_{[0,\infty)}$, where our applications lie. Consider a positive measure ν on $\mathcal{B}_{[0,\infty)}$. We say that ν is a Radon measure if $\nu(A) < \infty$, for every bounded set $A \in \mathcal{B}_{[0,\infty)}$. Let $\mathcal{M} := \mathcal{M}([0,\infty))$ denote the space of all Radon measures defined on $([0,\infty), \mathcal{B}_{[0,\infty)})$. We equip this space with the vague topology, i.e., $\nu_n \to \nu$ in $\mathcal{M}([0,\infty))$ whenever $\int \psi \, d\nu_n \to \int \psi \, d\nu$ for every continuous function $\psi : [0,\infty) \to \mathbb{R}$ with compact support. Consider the subsets of $\mathcal{M}$ defined by $\mathcal{M}_p := \{\nu \in \mathcal{M} : \nu(A) \in \mathbb{N}_0 \text{ for all } A \in \mathcal{B}_{[0,\infty)}\}$ and $\mathcal{M}_a := \{\nu \in \mathcal{M} : \nu \text{ is an atomic measure}\}$. A *random measure* M on $[0,\infty)$ is a random element of $\mathcal{M}$, *i.e.*, let $(\mathcal{X}, \mathcal{B}_{\mathcal{X}}, \mathbb{P})$ be a probability space, then any measurable $M : \mathcal{X} \to \mathcal{M}$ is a random measure on $[0,\infty)$. A *point process* N and an *atomic random measure* A are defined similarly as random elements on $\mathcal{M}_p$ and $\mathcal{M}_a$, respectively.

A point measure ν of $\mathcal{M}_p$ can be written as $\nu = \sum_{i=1}^{\infty} \delta_{x_i}$, where $x_1, x_2, \dots$ is a collection of not necessarily distinct points in $[0,\infty)$ and δ_{x_i} is the Dirac measure at x_i, *i.e.*, for every $A \in \mathcal{B}_{[0,\infty)}$, we have that $\delta_{x_i}(A) = 1$ if $x_i \in A$ and $\delta_{x_i}(A) = 0$, otherwise. The elements ν of $\mathcal{M}_a$ can be written as $\nu = \sum_{i=1}^{\infty} d_i \delta_{x_i}$, where $x_1, x_2, \dots \in [0,\infty)$ and $d_1, d_2, \dots \in [0,\infty)$. Hence, a point process can be written as $N = \sum_{i=1}^{\infty} \delta_{T_i}$ and an atomic random measure as $A = \sum_{i=1}^{\infty} D_i \delta_{T_i}$, where each T_i and D_i is a positive random variable defined on the probability space $(\mathcal{X}, \mathcal{B}_{\mathcal{X}}, \mathbb{P})$, for all $i \in \mathbb{N}$.

As pointed out in [Kar91], an atomic random measure can be seen as *marked point process*, $\sum_{i=1}^{\infty} \delta_{(T_i, D_i)}$, which is a point process on the higher dimensional space $[0,\infty) \times [0,\infty)$, where D_i

is called the *mark* associated to T_i. For this reason, from this point forward, we will refer to all atomic random measures as marked point processes.

A concrete example of a marked point process, which in particular will appear as the limit of the marked point processes, is the compound Poisson process which we define next.

Definition 2.1. Let $T_1, T_2, \ldots$ be an i.i.d. sequence of r.v. with common exponential distribution of mean $1/\theta$. Let $D_1, D_2, \ldots$ be another i.i.d. sequence of r.v., independent of the previous one, and with d.f. π. Given these sequences, for $J \in \mathcal{B}_{[0,\infty)}$, set

$$A(J) = \int \mathbf{1}_J \, d\left(\sum_{i=1}^{\infty} D_i \delta_{T_1+\ldots+T_i}\right).$$

Let $\mathcal{X}$ denote the space of all possible realisations of $T_1, T_2, \ldots$ and $D_1, D_2, \ldots$, equipped with a product σ-algebra and measure, then $A : \mathcal{X} \to \mathcal{M}_a([0,\infty))$ is a marked point process which we call a compound Poisson process of intensity θ and multiplicity d.f. π.

Remark 2.2. When $D_1, D_2, \ldots$ are integer valued positive random variables, π is completely defined by the values $\pi_k = \mathbb{P}(D_1 = k)$, for every $k \in \mathbb{N}_0$ and A is actually a point process. If $\pi_1 = 1$ and $\theta = 1$, then A is the standard Poisson process and, for every $t > 0$, the random variable $A([0,t))$ has a Poisson distribution of mean t.

Now, we will give a definition of convergence of random measures (for more details, see [Kal86]).

Definition 2.3. Let $(M_n)_{n\in\mathbb{N}} : \mathcal{X} \to \mathcal{M}$ be a sequence of random measures defined on a probability space $(\mathcal{X}, \mathcal{B}_{\mathcal{X}}, \mu)$ and let $M : Y \to \mathcal{M}$ be another random measure defined on a possibly distinct probability space $(Y, \mathcal{B}_Y, \nu)$. We say that M_n converges weakly to M if, for every bounded continuous function φ defined on $\mathcal{M}$, we have

$$\lim_{n\to\infty} \int \varphi d\mu \circ M_n^{-1} = \int \varphi d\nu \circ M^{-1}.$$

We write $M_n \overset{\mu}{\Longrightarrow} M$.

Checking the convergence of random measures using the definition is often quite hard, hence, it is useful to translate it into convergence in distribution of more tractable random variables or in terms of Laplace transforms. For that purpose, we let $\mathcal{S}$ denote the semi-ring of subsets of $\mathbb{R}_0^+$ whose elements are intervals of the type $[a,b)$, for $a, b \in \mathbb{R}_0^+$. Let $\mathcal{R}$ denote the ring generated by $\mathcal{S}$. Recall that for every $J \in \mathcal{R}$ there are $\varsigma \in \mathbb{N}$ and ς disjoint intervals $I_1, \ldots, I_\varsigma \in \mathcal{S}$ such that $J = \dot{\cup}_{i=1}^{\varsigma} I_j$. In order to fix notation, let $a_j, b_j \in \mathbb{R}_0^+$ be such that $I_j = [a_j, b_j) \in \mathcal{S}$.

Definition 2.4. Let Z be a non-negative, random variable with distribution F. For every $y \in \mathbb{R}_0^+$, the *Laplace transform* $\phi(y)$ of the distribution F is given by

$$\phi(y) := \mathbb{E}\left(\mathrm{e}^{-yZ}\right) = \int \mathrm{e}^{-yZ} d\mu_F,$$

where μ_F is the Lebesgue-Stieltjes probability measure associated to the distribution function F.

Definition 2.5. For a random measure M on $\mathbb{R}_0^+$ and ς disjoint intervals $I_1, I_2, \ldots, I_\varsigma \in \mathcal{S}$ and non-negative $y_1, y_2, \ldots, y_\varsigma$, we define the *joint Laplace transform* $\psi(y_1, y_2, \ldots, y_\varsigma)$ by

$$\psi_M(y_1, y_2, \ldots, y_\varsigma) = \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell M(I_\ell)}\right).$$

If M is a compound Poisson point process with intensity λ and multiplicity distribution π, then given ς disjoint intervals $I_1, I_2, \ldots, I_\varsigma \in \mathcal{S}$ and non-negative $y_1, y_2, \ldots, y_\varsigma$ we have:

$$\psi_M(y_1, y_2, \ldots, y_\varsigma) = \mathrm{e}^{-\lambda \sum_{\ell=1}^{\varsigma}(1-\phi(y_\ell))|I_\ell|},$$

where $\phi(y)$ is the Laplace transform of the multiplicity distribution π.

Remark 2.6. By [Kal86, Theorem 4.2], the sequence of random measures $(M_n)_{n\in\mathbb{N}}$ converges weakly to the random measure M iff the sequence of vector r.v. $(M_n(J_1), \ldots, M_n(J_\varsigma))$ converges in distribution to $(M(J_1), \ldots, M(J_\varsigma))$, for every $\varsigma \in \mathbb{N}$ and all disjoint $J_1, \ldots, J_\varsigma \in \mathcal{S}$ such that $M(\partial J_\ell) = 0$ a.s., for $\ell = 1, \ldots, \varsigma$, which will be the case if the respective joint Laplace transforms $\psi_{M_n}(y_1, y_2, \ldots, y_\varsigma)$ converge to the joint Laplace transform $\psi_M(y_1, y_2, \ldots, y_\varsigma)$, for all $y_1, \ldots, y_\varsigma \in [0, \infty)$.

2.2. **Marked Point Processes of Rare Events.** Before we give the formal definition of Marked Point Processes of Rare Events, we need to introduce some notation and definitions that will also be useful to understand the conditions that we will introduce in order to prove their weak convergence.

In what follows, for every $A \in \mathcal{B}$, we denote the complement of A as $A^c := \mathcal{Y} \setminus A$.

Given a set of thresholds $u_{n,i}$, for each n, i and $j \in \mathbb{N}_0$ with $j < Hn - i$, we set

$$U^{(0)}_{j,n,i} := \{X_i > u_{n,i}\}$$

$$Q^{(0)}_{j,n,i} := U^{(0)}_{j,n,i} \cap \bigcap_{\ell=1}^{j} (U^{(0)}_{j,n,i+\ell})^c = \{X_i > u_{n,i}, X_{i+1} \le u_{n,i+1}, ..., X_{i+j} \le u_{n,i+j}\}$$

and define the following events, for $\kappa \in \mathbb{N}$:

$$U^{(\kappa)}_{j,n,i} := U^{(\kappa-1)}_{j,n,i} \setminus Q^{(\kappa-1)}_{j,n,i} = U^{(\kappa-1)}_{j,n,i} \cap \bigcup_{\ell=1}^{j} U^{(\kappa-1)}_{j,n,i+\ell}$$

$$Q^{(\kappa)}_{j,n,i} := U^{(\kappa)}_{j,n,i} \cap \bigcap_{\ell=1}^{j} (U^{(\kappa)}_{j,n,i+\ell})^c.$$

If $j = 0$ then $Q^{(0)}_{0,n,i} = U^{(0)}_{0,n,i} = \{X_i > u_{n,i}\}$ and $Q^{(\kappa)}_{0,n,i} = U^{(\kappa)}_{0,n,i} = \emptyset$ for $\kappa \in \mathbb{N}$.

For $j \ge Hn - i$, we set $Q^{(\kappa)}_{j,n,i} = U^{(\kappa)}_{j,n,i} = \emptyset$ for all $\kappa \in \mathbb{N}_0$.

Also, let $U^{(\infty)}_{j,n,i} := \bigcap_{\kappa=0}^{\infty} U^{(\kappa)}_{j,n,i}$. Note that $Q^{(\kappa)}_{j,n,i} = U^{(\kappa)}_{j,n,i} \setminus U^{(\kappa+1)}_{j,n,i}$ for $\kappa \in \mathbb{N}_0$ and, therefore,

$$U^{(0)}_{j,n,i} = \bigcup_{\kappa=0}^{\infty} Q^{(\kappa)}_{j,n,i} \cup U^{(\infty)}_{j,n,i}.$$

Remark 2.7. The points in $U^{(\kappa)}_{j,n,i}$ are points whose orbit represents a cluster of size at least $\kappa + 1$ (that is, in their orbit there are at least $\kappa + 1$ exceedances separated by at most j time steps between subsequent ones), since there will be points in each $U^{(\kappa')}_{j,n,i}$ with k' taking values between κ and 0. On the other hand, points in $Q^{(\kappa)}_{j,n,i}$ are points whose orbit represents a cluster of size $\kappa+1$ exactly (that is, in their orbit there are exactly $\kappa + 1$ exceedances separated by at most j time

steps between subsequent ones). The underlying method to identify clusters we are using here is called *runs declustering scheme*, which sets a fixed $j \in \mathbb{N}$ as the maximum waiting time between the occurrence of two extreme events on the same cluster, so that any rare events belong to the same cluster when they are separated by at most $j-1$ non-extreme observations.

For each $i \in \mathbb{N}_0$ and $n \in \mathbb{N}$, let $R_{j,n,i} := \min\{r \in \mathbb{N} : Q^{(0)}_{j,n,i} \cap Q^{(0)}_{j,n,i+r} \neq \emptyset\}$. We assume that there exists $q \in \mathbb{N}_0$ such that:

$$q = \min\left\{ j \in \mathbb{N}_0 : \lim_{n\to\infty} \min_{i \le n} \{R_{j,n,i}\} = \infty \right\}. \tag{2.6}$$

Note that one can view q as the largest of the periods of the underlying periodic phenomena present in the stochastic process, which, in the dynamical context, is related with the periodicity of the maximal set of points where the observable achieves the global maximum.

When $q = 0$ then $Q^{(0)}_{0,n,i}$ corresponds to an exceedance of the threshold $u_{n,i}$ and we expect no clustering of exceedances.

When $q > 0$, heuristically one can think that there exists an underlying periodic phenomenon creating short recurrence, *i.e.*, clustering of exceedances, when exceedances occur separated by at most q units of time then they belong to the same cluster. Hence, the sets $Q^{(0)}_{q,n,i}$ correspond to the occurrence of exceedances that escape the periodic phenomenon and are not followed by another exceedance in the same cluster. We will refer to the occurrence of $Q^{(0)}_{q,n,i}$ as the occurrence of an escape at time i, whenever $q > 0$.

Given an interval $I \in \mathcal{S}$, $x \in \mathcal{X}$ and $u_{n,i} \in \mathbb{R}$, we define

$$N_{n,I}(x) := \sum_{i \in I \cap \mathbb{N}_0} \mathbf{1}_{Q^{(0)}_{q,n,i}}(x).$$

Let $i_1(x) < i_2(x) < \ldots < i_{N_{n,I}(x)}(x)$ denote the times at which the orbit of x entered $Q^{(0)}_{q,n,i}$ in I. We now define the cluster periods: for every $k = 1, \ldots, N_{n,I}(x) - 1$ let $I_k(x) = (i_k(x), i_{k+1}(x)]$ and set $I_0(x) = [\min I, i_1(x)]$ and $I_{N_{n,I}(x)}(x) = (i_{N_{n,I}(x)}(x), \sup I]$.

In order to define the marks for each cluster we consider the following mark functions that depend on the levels $u_{n,i}$ and on the random variables in a certain time frame $I \in \mathcal{S}$:

$$m_n(I) := \begin{cases} \sum_{i \in I \cap \mathbb{N}_0} (X_i - u_{n,i})_+ & \text{AOT case} \\ \max_{i \in I \cap \mathbb{N}_0} \{(X_i - u_{n,i})_+\} & \text{POT case} \\ \sum_{i \in I \cap \mathbb{N}_0} \mathbf{1}_{X_i > u_{n,i}} & \text{REPP case,} \end{cases} \tag{2.7}$$

where $(y)_+ = \max\{y, 0\}$ and when $I \cap \mathbb{N}_0 \neq \emptyset$. Also set $m_n(I) := 0$ when $I \cap \mathbb{N}_0 = \emptyset$.

Finally, we set

$$\mathscr{A}_n(I)(x) := \sum_{k=0}^{N_{n,I}(x)} m_n(I_k(x)).$$

In order to avoid degeneracy problems in the definition of the marked point processes we need to rescale time by the factor

$$v_n := n / F^*_{1,n}$$

so that the expected average number of exceedances of the levels $u_{n,i}$ for $i = 0, \dots, n$ in each time frame considered is kept 'constant' as $n \to \infty$. Recall that the levels $u_{n,i}$ satisfy 2.5, and therefore $v_n \sim \frac{n}{\tau}$, where we use the notation $A(n) \sim B(n)$, when $\lim_{n\to\infty} \frac{A(n)}{B(n)} = 1$. Hence, we introduce the following notation. For $I = [a, b) \in \mathcal{S}$ and $\alpha \in \mathbb{R}$, we denote $\alpha I := [\alpha a, \alpha b)$ and $I+\alpha := [a+\alpha, b+\alpha)$. Similarly, for $J \in \mathcal{R}$, such that $J = J_1 \dot\cup \dots \dot\cup J_k$, define $\alpha J := \alpha J_1 \dot\cup \cdots \dot\cup \alpha J_k$ and $J + \alpha := (J_1 + \alpha) \dot\cup \cdots \dot\cup (J_k + \alpha)$.

Definition 2.8. We define the *marked rare event point process* (MREPP) by setting for every $J \in \mathcal{R}$, with $J = J_1 \dot\cup \dots \dot\cup J_k$, where $J_i \in \mathcal{S}$ for all $i = 1, \dots, k$,

$$A_n(J) := \sum_{i=1}^{k} \mathscr{A}_n(v_n J_i). \tag{2.8}$$

When m_n given in (2.7) is as in the AOT case, then the MREPP A_n computes the sum of all excesses over the threshold u_n and, in such case, we will refer to A as being an *area over threshold* or AOT MREPP. Observe that in this case A_n does not rely on the definition of clusters but takes into account each exceedance, since we may write:

$$A_n(J) = \sum_{i \in v_n J \cap \mathbb{N}_0} (X_i - u_{n,i})_+.$$

When m_n given in (2.7) is as in the POT case, then the MREPP A_n computes the sum of the largest excess (peak) over the threshold $u_{n,i}$ within each cluster and, in such case, we will refer to A_n as being a *peaks over threshold* or POT MREPP.

When m_n given in (2.7) is as in the REPP case, then the MREPP A_n is actually a point process that counts the number of exceedances of $u_{n,i}$ and, in such case, we will refer to A_n as being a *rare events point process* or REPP, as it was referred in [FFT13]. Also in this case A_n does not rely on the definition of clusters but takes into account each exceedance, since we may write:

$$A_n(J) = \sum_{i \in v_n J \cap \mathbb{N}_0} \mathbf{1}_{X_i > u_{n,i}}.$$

If $q = 0$ then the AOT MREPP and the POT MREPP coincide and both compute the sum of all excesses over the threshold $u_{n,i}$. In such situation we say that A_n is an *excesses over threshold* (EOT) MREPP.

Next, we will introduce the dependence conditions $Д_q(u_{n,i})^*$ and $Д'_q(u_{n,i})^*$, which are the analogous of conditions $Д_p(u_n)$ and $Д'_p(u_n)$ considered in [FFT15], but designed to establish the convergence of MREPP (AOT, POT or REPP), which allow us to state our main result. Before we do that, we need to introduce some additional notation and definitions.

For $x \geq 0$ and $\kappa \in \mathbb{N}_0$, we define the following events:

$$R_{n,i}^{(\kappa)}(x) := Q_{q,n,i}^{(\kappa)} \cap \{m_n(I_\kappa) > x\}$$

$$B_{n,i}(x) := \bigcup_{\kappa=0}^{\infty} R_{n,i}^{(\kappa)}(x) \cup U_{q,n,i}^{(\infty)}$$

$$A_{n,i}(x) := B_{n,i}(x) \cap \bigcap_{\ell=1}^{q} (B_{n,i+\ell}(x))^c$$

where $I_\kappa = [i, i_{\kappa,\kappa}+1)$ and $i_{\kappa,j}$ denotes the times at which the orbit of the considered point entered $Q_{q,n,i}^{(\kappa-j)}$, with $i = i_{\kappa,0} < i_{\kappa,1} < i_{\kappa,2} < \ldots < i_{\kappa,j} < \ldots < i_{\kappa,\kappa}$ (see Remark 2.7). Note that one can view $B_{n,i}(x)$ as the set of points in $U_{j,n,i}^{(0)}$ whose corresponding cluster as a mark greater than x, while $R_{n,i}^{(\kappa)}(x)$ are the points in $U_{j,n,i}^{(0)}$ whose corresponding cluster has size $\kappa+1$ and a mark greater than x.

In particular, for $x = 0$ we have

$$\begin{aligned} R_{n,i}^{(\kappa)}(0) &= Q_{q,n,i}^{(\kappa)} \\ B_{n,i}(0) &= \bigcup_{\kappa=0}^{\infty} Q_{q,n,i}^{(\kappa)} \cup U_{q,n,i}^{(\infty)} = U_{q,n,i}^{(0)} \\ A_{n,i}(0) &= U_{q,n,i}^{(0)} \cap \bigcap_{\ell=1}^{q} (U_{q,n,i+\ell}^{(0)})^c = Q_{q,n,i}^{(0)} \end{aligned}$$

and, if $q = 0$,

$$\begin{aligned} &R_{n,i}^{(0)}(x) = \{X_i > u_{n,i}, m_n([i,i+1)) > x\}, \quad R_{n,i}^{(\kappa)}(x) = \emptyset \text{ for } \kappa \in \mathbb{N} \\ &A_{n,i}(x) = B_{n,i}(x) = R_{n,i}^{(0)}(x). \end{aligned}$$

Condition ($Д_q(u_{n,i})^*$)**.** We say that $Д_q(u_{n,i})^*$ holds for the sequence $X_0, X_1, X_2, \ldots$ if for $t, n \in \mathbb{N}$, $i = 0, \ldots, Hn-1$, for $x_1, \ldots, x_\varsigma \geq 0$ and any $J = \cup_{i=2}^{\varsigma} I_j \in \mathcal{R}$ with $\inf\{x : x \in J\} \geqslant i+t$,

$$\left| \mathbb{P}\left(A_{n,i}(x_1) \cap \bigcap_{j=2}^{\varsigma} \{\mathscr{A}_n(I_j) \leq x_j\} \right) - \mathbb{P}\left(A_{n,i}(x_1)\right) \mathbb{P}\left(\bigcap_{j=2}^{\varsigma} \{\mathscr{A}_n(I_j) \leq x_j\} \right) \right| \leq \gamma_i(n,t),$$

where for each n and each i we have that $\gamma_i(n,t)$ is nonincreasing in t and there exists a sequence $t_n^* = o(n)$ such that $t_n^* \bar{F}_{n,\max}(H) \to 0$ and $n\gamma_i(n, t_n^*) \to 0$ when $n \to \infty$.

Note that the main advantage of this mixing condition when compared with condition $\Delta(u_n)$ used by Leadbetter in [Lea91] or any other similar such condition available in the literature is that it follows easily from sufficiently fast decay of correlations and therefore is particularly useful when applied to stochastic processes arising from dynamical systems.

For $q \in \mathbb{N}_0$ given by (2.6), consider the sequence $(t_n^*)_{n\in\mathbb{N}}$, given by condition $Д_q(u_{n,i})^*$ and let $(k_n)_{n\in\mathbb{N}}$ be another sequence of integers such that

$$k_n \to \infty \quad \text{and} \quad k_n t_n^* \bar{F}_{n,\max}(H) \to 0 \tag{2.9}$$

as $n \to \infty$ for every $H \in \mathbb{N}$.

Let us give a brief description of the blocking argument and postpone the precise construction of the blocks to Section 3.1. We split the data into k_n blocks separated by time gaps of size larger than t_n^*, where we simply disregard the observations in the corresponding time frame. In the stationary case, the blocks have the same size and the expected number of exceedances within each block is $\sim \tau/k_n$. Here, the blocks may have different sizes, denoted by $\ell_{H,n,1}, \ldots, \ell_{H,n,k_n}$, but, as in [FFV17], these are chosen so that the expected number of exceedances is again $\sim \tau/k_n$. Also, for $i = 1, \ldots, k_n$, let $\mathcal{L}_{H,n,i} = \sum_{j=1}^{i} \ell_{H,n,j}$ and $\mathcal{L}_{H,n,0} = 0$.

Note that gaps need to be big enough so that they are larger than t_n^* but they also need to be sufficiently small so that the information disregarded does not compromise the computations. This is achieved by choosing the number of blocks, which correspond to the sequence $(k_n)_{n\in\mathbb{N}}$, diverging but slowly enough so that the weight of the gaps is negligible when compared to that of the true blocks.

In order to guarantee the existence of a distributional limit one needs to impose some restrictions on the speed of recurrence.

Condition ($Д'_q(u_{n,i})^*$)**.** We say that $Д'_q(u_{n,i})^*$ holds for the sequence $X_0, X_1, X_2, \ldots$ if there exists a sequence $(k_n)_{n\in\mathbb{N}}$ satisfying (2.9) and such that, for every $H \in \mathbb{N}$,

$$\lim_{n\to\infty} \sum_{i=1}^{k_n} \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \sum_{r>j}^{\mathcal{L}_{H,n,i}-1} \mathbb{P}\left(Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}\right) = 0 \tag{2.10}$$

$$\text{and} \quad \lim_{n\to\infty} \sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1} \sum_{r>j}^{Hn-1} \mathbb{P}\left(Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}\right) = 0 \tag{2.11}$$

Condition $Д'_q(u_{n,i})^*$ precludes the occurrence of clustering of escapes (or exceedances, when $q = 0$).

Remark 2.9. Note that condition $Д'_q(u_{n,i})^*$ is an adjustment of a similar condition $Д'_p(u_n)$ in [FFT15] in the stationary setting, which is similar to condition $D^{(p+1)}(u_n)$ in the formulation of [CHM91, Equation (1.2)], although slightly weaker.

When $q = 0$, observe that $Д'_q(u_{n,i})^*$ is very similar to $D'(u_{n,i})$ from Hüsler, which prevents clustering of exceedances, just as $D'(u_n)$ introduced by Leadbetter did in the stationary setting.

When $q > 0$, we have clustering of exceedances, *i.e.*, the exceedances have a tendency to appear aggregated in groups (called clusters). One of the main ideas in [FFT12] that we use here is that the events $Q^{(0)}_{q,n,i}$ play a key role in determining the limiting EVL and in identifying the clusters. In fact, when $Д'_q(u_{n,i})^*$ holds we have that every cluster ends with an entrance in $Q^{(0)}_{q,n,i}$, which means that the inter cluster exceedances must appear separated at most by q units of time.

In this approach, condition $Д'_q(u_{n,i})^*$ plays a prominent role. In particular, note that if condition $Д'_q(u_{n,i})^*$ holds for some particular $q = q_0 \in \mathbb{N}_0$, then it holds for all $q \geq q_0$, and so (2.6) is indeed the natural candidate to try to show the validity of $Д'_q(u_{n,i})^*$.

Now, we give a way of defining the Extremal Index (EI) using the sets $Q^{(0)}_{q,n,i}$. For $q \in \mathbb{N}_0$ given by (2.6), we also assume that there exists $0 \leq \theta \leq 1$, which will be referred to as the EI, such that

$$\lim_{n\to\infty} \max_{i=1,\ldots,k_n} \left\{ \left| \theta k_n \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \bar{F}_j(u_{n,j}) - k_n \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \mathbb{P}\left(Q^{(0)}_{q,n,j}\right) \right| \right\} = 0. \tag{2.12}$$

Remark 2.10. If the process is stationary, then $\ell_{H,n,i} \sim Hn/k_n$ and the previous condition becomes

$$\lim_{n\to\infty} Hn \left| \theta \mathbb{P}(X > u_{n,0}) - \mathbb{P}\left(Q^{(0)}_{q,n,0}\right) \right| = 0$$

so that, using the usual hypothesis $n\mathbb{P}(X > u_{n,0}) \to \tau$, we have

$$\theta = \lim_{n\to\infty} \frac{\mathbb{P}\left(Q^{(0)}_{q,n,0}\right)}{\mathbb{P}(X > u_{n,0})}$$

and θ is given by the usual definition, as in [FFMa18]

Moreover, we assume the existence of normalising factors $a_{n,j}$ for every $j = 0, 1, \ldots, Hn-1$ and $n \in \mathbb{N}$, and a probability distribution π such that, for every $H \in \mathbb{N}$ and $x \geq 0$,

$$\lim_{n\to\infty} \max_{j=0,1,\ldots,Hn-1} \left\{ \left| \frac{\mathbb{P}(A_{n,j}(x/a_{n,j}))}{\mathbb{P}\left(Q^{(0)}_{q,n,j}\right)} - (1-\pi(x)) \right| \right\} = 0 \tag{2.13}$$

and in this way obtain a formula to compute the multiplicity distribution of the limiting compound Poisson process.

Finally, assuming that both $Д_q(u_{n,i})^*$ and $Д'_q(u_{n,i})^*$ hold, we give a technical condition which imposes a sufficiently fast decay of the probability of having very long clusters. We will call it $ULC_q(u_{n,i})$ that stands for 'Unlikely Long Clusters'. Of course this condition is trivially satisfied when there is no clustering.

Condition ($ULC_q(u_{n,i})$)**.** We say that condition $ULC_q(u_{n,i})$ holds if, for all $H \in \mathbb{N}$ and $y > 0$,

$$\lim_{n\to\infty} \sum_{i=1}^{k_n} \int_0^\infty \mathrm{e}^{-yx} \delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}}(x/a_n)dx = 0,$$

$$\lim_{n\to\infty} \int_0^\infty \mathrm{e}^{-x} \delta_{n,\mathcal{L}_{H,n,k_n},Hn-\mathcal{L}_{H,n,k_n}}(x/a_n)dx = 0,$$

$$\text{and} \quad \lim_{n\to\infty} \sum_{i=1}^{k_n} \int_0^\infty \mathrm{e}^{-yx} \delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}-t_{H,n,i}}(x/a_n)dx = 0$$

where a_n is such that $\mathbb{P}(A_{n,j}(x/a_{n,j})) = \mathbb{P}(A_{n,j}(x/a_n))$ for $a_{n,j}$ is as in (2.13), $\delta_{n,s,\ell}(x) := 0$ for $q = 0$ and, for $q > 0$,

$$\delta_{n,s,\ell}(x) := \sum_{\kappa=1}^{\lfloor \ell/q \rfloor} \sum_{j=s+\ell-\kappa q}^{s+\ell-1} \mathbb{P}\left(R^{(\kappa)}_{n,j}(x)\right) + \sum_{j=s}^{s+\ell-1} \sum_{\kappa > \lfloor \ell/q \rfloor} \mathbb{P}\left(R^{(\kappa)}_{n,j}(x)\right) + \sum_{j=1}^{q} \mathbb{P}(B_{n,s+\ell-j}(x)) \tag{2.14}$$

is an integrable function in $\mathbb{R}^+$ for n sufficiently large.

Note that, by definition, condition $ULC_0(u_{n,i})$ always holds. Note also that $\delta_{n,s,\ell}(x) \leq \delta_{n,s',\ell'}(x)$ if $s + \ell = s' + \ell'$ and $\ell \leq \ell'$. In particular, if $ULC_q(u_{n,i})$ holds then, for all $H \in \mathbb{N}$ and $y > 0$,

$$\lim_{n\to\infty} \sum_{i=1}^{k_n} \int_0^\infty \mathrm{e}^{-yx} \delta_{n,\mathcal{L}_{H,n,i}-t_{H,n,i},t_{H,n,i}}(x/a_n)dx = 0$$

We are now ready to state the main convergence result:

Theorem 2.A. *Let $X_0, X_1, \dots$ be given by* (2.1) *and $u_{n,i}$ be real-valued boundaries satisfying* (2.2) *and* (2.3)*. Assume that $Д_q(u_{n,i})^*$, $Д'_q(u_{n,i})^*$ and $ULC_q(u_{n,i})^*$ hold, for some $q \in \mathbb{N}_0$. Assume the existence of θ satisfying* (2.12) *and a normalising sequence $(a_n)_{n\in\mathbb{N}}$ such that $\mathbb{P}(A_{n,j}(x/a_{n,j})) = \mathbb{P}(A_{n,j}(x/a_n))$ for any $j = 0, 1, \dots, Hn-1$, where $a_{n,j}$ are normalising factors such that* (2.13) *holds for some probability distribution π. Then, the MREPP a_nA_n, where A_n is given by Definition* 2.8 *for any of the 3 mark functions considered in* (2.7)*, converges in distribution to a compound Poisson process A with intensity θ and multiplicity d.f. π.*

Remark 2.11. If the normalising factors $a_{n,j}$ don't depend on j, then we can naturally choose $a_n = a_{n,j}$ for every $n \in \mathbb{N}$.

Remark 2.12. What is essential, about the mark function m_u considered in (2.7) to define the respective MREPP, is that it satisfies the following assumptions:

(1) $m_n(I) \geq 0$ and $m_n(\emptyset) = 0$
(2) $m_n(I) \leq m_n(J)$ if $I \subset J$
(3) $m_n(I) = m_n(J)$ if $X_i \leq u_{n,i}, \forall i \in (I \setminus J) \cap \mathbb{N}_0$

Note that, in particular, we must have $m_n(I) = 0$ if $X_i \leq u_{n,i}, \forall i \in I \cap \mathbb{N}_0$.

As long as the above assumptions hold then the conclusion of Theorem 2.A holds for the MREPP defined from such a mark function m_n satisfying the three assumptions just enumerated.

3. Convergence of marked rare events point processes

This section is dedicated to the proof of Theorem 2.A, whose argument follows the same thread as the one in the proof of [FFMa18, Theorem 2.A.]

3.1. The construction of the blocks.

The construction of the blocks is designed so that the expected number of exceedances in each block is the same. We follow closely the construction in [FFV17], which was inspired in [Hüs83, Hüs86].

For each $H, n \in \mathbb{N}$ we split the random variables $X_0, \dots, X_{Hn-1}$ into k_n initial blocks, where k_n is given by (2.9), of sizes $\ell_{H,n,1}, \dots, \ell_{H,n,k_n}$ defined in the following way. Let as before $\mathcal{L}_{H,n,i} = \sum_{j=1}^{i} \ell_{H,n,j}$ and $\mathcal{L}_{H,n,0} = 0$. Assume that $\ell_{H,n,1}, \dots, \ell_{H,n,i-1}$ are already defined. Take $\ell_{H,n,i}$ to be the largest integer such that

$$\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i-1}+\ell_{H,n,i}-1} \bar{F}(u_{n,j}) \leq \frac{F^*_{H,n}}{k_n}.$$

The final working blocks are obtained by disregarding the last observations of each initial block, which will create a time gap between each final block. The size of the time gaps must be balanced in order to have at least a size t_n^* but such that its weight on the average number of exceedances is negligible when compared to that of the final blocks. For that purpose we define

$$\varepsilon(H,n) := (t_n^* + 1)\bar{F}_{n,\max}(H)\frac{k_n}{F^*_{H,n}}.$$

Note that by (2.3) and (2.9), it follows immediately that $\lim_{n\to\infty}\varepsilon(H,n)=0$. Now, for each $i=1,\dots,k_n$ let $t_{H,n,i}$ be the largest integer such that

$$\sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})\le\varepsilon(H,n)\frac{F^*_{H,n}}{k_n}.$$

Hence, the final working blocks correspond to the observations within the time frame $\mathcal{L}_{H,n,i-1},\dots,\mathcal{L}_{H,n,i}-t_{H,n,i}-1$, while the time gaps correspond to the observations in the time frame $\mathcal{L}_{H,n,i}-t_{H,n,i},\dots,\mathcal{L}_{H,n,i}-1$, for all $i=1,\dots,k_n$.

Note that $t^*_n<t_{H,n,i}<\ell_{H,n,i}$, for each $i=1,\dots,k_n$. The second inequality is trivial. For the first inequality note that by definition of $t_{H,n,i}$ we have

$$\varepsilon(H,n)\frac{F^*_{H,n}}{k_n}<\sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}-1}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})\le(t_{H,n,i}+1)\bar{F}_{n,\max}(H).$$

The first inequality follows easily now by definition of $\varepsilon(H,n)$.

Also, note that, by choice of $\ell_{H,n,i}$ we have

$$\frac{F^*_{H,n}}{k_n}\le\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})+\bar{F}(u_{n,\mathcal{L}_{H,n,i}})\le\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})+\bar{F}_{n,\max}(H)$$

and then it follows that

$$\frac{F^*_{H,n}}{k_n}-\bar{F}_{n,\max}(H)\le\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})\le\frac{F^*_{H,n}}{k_n}. \tag{3.1}$$

From the first inequality we get

$$F^*_{H,n}-k_n\bar{F}_{n,\max}(H)\le\sum_{i=1}^{k_n}\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})$$

which implies that

$$\sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1}\bar{F}(u_{n,j})=F^*_{H,n}-\sum_{i=1}^{k_n}\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})\le k_n\bar{F}_{n,\max}(H) \tag{3.2}$$

which goes to 0 as $n\to\infty$ by (2.9).

Let $\mathbb{A}:=(A_0,A_1,\dots)$ be a sequence of events such that $A_i\in\mathcal{T}_i^{-1}\mathcal{B}$. For some $s,\ell\in\mathbb{N}_0$, we define

$$\mathscr{W}_{s,\ell}(\mathbb{A})=\bigcap_{i=s}^{s+\ell-1}A_i^c, \tag{3.3}$$

which forbids the occurrence of A_i during the time interval between s and $s+\ell-1$.

Proposition 3.1. *Given events $B_0, B_1, \ldots \in \mathcal{B}$, let $r, q, s, \ell \in \mathbb{N}$ be such that $q < n$ and define $\mathbb{B} = (B_0, B_1, \ldots)$, $A_r = B_r \setminus \bigcup_{j=1}^{q} B_{r+j}$ and $\mathbb{A} = (A_0, A_1, \ldots)$. Then*

$$|\mathbb{P}(\mathscr{W}_{s,\ell}(\mathbb{B})) - \mathbb{P}(\mathscr{W}_{s,\ell}(\mathbb{A}))| \leq \sum_{j=1}^{q} \mathbb{P}\left(\mathscr{W}_{s,\ell}(\mathbb{A}) \cap (B_{s+\ell-j} \setminus A_{s+\ell-j})\right).$$

Proof. Since $A_r \subset B_r$, then clearly $\mathscr{W}_{s,\ell}(\mathbb{B}) \subset \mathscr{W}_{s,\ell}(\mathbb{A})$. Hence, we have to estimate the probability of $\mathscr{W}_{s,\ell}(\mathbb{A}) \setminus \mathscr{W}_{s,\ell}(\mathbb{B})$.

Let $x \in \mathscr{W}_{s,\ell}(\mathbb{A}) \setminus \mathscr{W}_{s,\ell}(\mathbb{B})$. We will see that there exists $j \in \{1, \ldots, q\}$ such that $x \in B_{s+\ell-j}$. In fact, suppose that no such j exists. Then let $\kappa = \max\{i \in \{s, \ldots, s+\ell-1\} : x \in B_i\}$. Then, clearly, $\kappa < s + \ell - q$. Hence, if $x \notin B_j$, for all $i = \kappa + 1, \ldots, s + \ell - 1$, then we must have that $x \in A_\kappa$ by definition of A. But this contradicts the fact that $x \in \mathscr{W}_{s,\ell}(\mathbb{A})$. Consequently, we have that there exists $j \in \{1, \ldots, q\}$ such that $x \in B_{s+\ell-j}$ and since $x \in \mathscr{W}_{s,\ell}(\mathbb{A})$ then we can actually write $x \in B_{s+\ell-j} \setminus A_{s+\ell-j}$.

This means that $\mathscr{W}_{s,\ell}(\mathbb{A}) \setminus \mathscr{W}_{s,\ell}(\mathbb{B}) \subset \bigcup_{j=1}^{q}(B_{s+\ell-j} \setminus A_{s+\ell-j}) \cap \mathscr{W}_{s,\ell}(\mathbb{A})$ and then

$$\begin{aligned}\left|\mathbb{P}(\mathscr{W}_{s,\ell}(\mathbb{B})) - \mathbb{P}(\mathscr{W}_{s,\ell}(\mathbb{A}))\right| &= \mathbb{P}(\mathscr{W}_{s,\ell}(\mathbb{A}) \setminus \mathscr{W}_{s,\ell}(\mathbb{B})) \\ &\leq \mathbb{P}\left(\bigcup_{j=1}^{q}(B_{s+\ell-j} \setminus A_{s+\ell-j}) \cap \mathscr{W}_{s,\ell}(\mathbb{A})\right) \leq \sum_{j=1}^{q} \mathbb{P}\left(\mathscr{W}_{s,\ell}(\mathbb{A}) \cap (B_{s+\ell-j} \setminus A_{s+\ell-j})\right),\end{aligned}$$

as required. □

Applying this proposition to $B_i = B_{n,i}(x)$, we have the following lemma, which says that the probability of not entering $B_{n,i}(x)$ can be approximated by the probability of not entering $A_{n,i}(x)$ during the same period of time.

Lemma 3.2. *For any $s, \ell \in \mathbb{N}$ and $x \geq 0$ we have*

$$\left|\mathbb{P}\big(\mathscr{W}_{s,\ell}(B_{n,i}(x))\big) - \mathbb{P}\big(\mathscr{W}_{s,\ell}(A_{n,i}(x))\big)\right| \leq \sum_{i=1}^{q} \mathbb{P}(B_{n,s+\ell-i}(x))$$

Next we give an approximation for the probability of not entering $A_{n,i}(x)$ during a certain period of time.

Lemma 3.3. *For any $s, \ell \in \mathbb{N}$ and $x \geq 0$ we have*

$$\left|\mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x))) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right)\right| \leq \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right)$$

Proof. Since $(\mathscr{W}_{s,\ell}(A_{n,i}(x)))^c = \cup_{i=s}^{s+\ell-1} A_{n,i}(x)$ it is clear that

$$\left|1 - \mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x))) - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right| \leq \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}(A_{n,j}(x) \cap A_{n,r}(x))$$

If $q > 0$, the result follows by the fact that $A_{n,r}(x) \subset \{X_r > u_{n,r}\}$ and the fact that the occurrence of both $A_{n,j}(x)$ and $A_{n,r}(x)$ implies an escape, *i.e.*, the occurrence of $Q_{q,n,j_1}^{(0)}$ for some $j \leq j_1 < r$

(otherwise, the occurrence of $A_{n,r}(x)$ and therefore of $B_{n,r}(x)$ would imply the occurrence of $B_{n,r_1}(x)$ for some $j+1 \le r_1 \le j+q$ which would contradict the occurrence of $A_{n,j}(x)$).

If $q=0$, the result follows immediately since $A_{n,i}(x) \subset \{X_i > u_{n,i}\} = Q^{(0)}_{0,n,i}$.

□

The next lemma gives an error bound for the approximation of the probability of the process $\mathscr{A}_n([s,s+\ell))$ not exceeding x by the probability of not entering in $B_{n,i}(x)$ during the period $[s,s+\ell)$. In what follows, we use the notation $\mathscr{A}^{s+\ell}_{n,s} := \mathscr{A}_n([s,s+\ell))$.

Lemma 3.4. *For any $s,\ell \in \mathbb{N}$ and $x \ge 0$ we have*

$$\left|\mathbb{P}(\mathscr{A}^{s+\ell}_{n,s} \le x) - \mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x)))\right| \le \sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}\right)$$
$$+ \sum_{\kappa=1}^{\lfloor \ell/q \rfloor}\sum_{i=s+\ell-\kappa q}^{s+\ell-1} \mathbb{P}\left(R^{(\kappa)}_{n,i}(x)\right) + \sum_{i=s}^{s+\ell-1}\sum_{\kappa>\lfloor \ell/q \rfloor} \mathbb{P}\left(R^{(\kappa)}_{n,i}(x)\right)$$

if $q>0$, and in case $q=0$ we have

$$\left|\mathbb{P}(\mathscr{A}^{s+\ell}_{n,s} \le x) - \mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x)))\right| \le \sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1} \mathbb{P}(X_j > u_{n,j}, X_r > u_{n,r}).$$

Proof. If $q>0$, we start by observing that

$$A^*_{n,s,\ell}(x) := \left\{\mathscr{A}^{s+\ell}_{n,s} \le x\right\} \cap (\mathscr{W}_{s,\ell}(B_{n,i}(x)))^c \subset$$
$$\bigcup_{i=s+\ell-q}^{s+\ell-1} R^{(1)}_{n,i}(x) \cup \bigcup_{i=s+\ell-2q}^{s+\ell-1} R^{(2)}_{n,i}(x) \cup \ldots \cup \bigcup_{i=s+\ell-\lfloor \ell/q \rfloor q}^{s+\ell-1} R^{(\lfloor \ell/q \rfloor)}_{n,i}(x) \cup \bigcup_{i=s}^{s+\ell-1}\bigcup_{\kappa>\lfloor \ell/q \rfloor} R^{(\kappa)}_{n,i}(x)$$

since $\bigcup_{i=s}^{s+\ell-\kappa q-1} R^{(\kappa)}_{n,i}(x) \subset \left\{\mathscr{A}^{s+\ell}_{n,s} > x\right\}$ for any $\kappa \le \lfloor \ell/q \rfloor$. So,

$$\mathbb{P}(A^*_{n,s,\ell}(x)) \le \sum_{\kappa=1}^{\lfloor \ell/q \rfloor}\sum_{i=s+\ell-\kappa q}^{s+\ell-1} \mathbb{P}\left(R^{(\kappa)}_{n,i}(x)\right) + \sum_{i=s}^{s+\ell-1}\sum_{\kappa>\lfloor \ell/q \rfloor} \mathbb{P}\left(R^{(\kappa)}_{n,i}(x)\right).$$

Now, we note that

$$B^*_{n,s,\ell}(x) := \left\{\mathscr{A}^{s+\ell}_{u,s} > x\right\} \cap \mathscr{W}_{s,\ell}(B_{n,i}(x)) \subset \bigcup_{j=s}^{s+\ell-1}\bigcup_{r=j+1}^{s+\ell-1} Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}.$$

This is because no entrance in $A_{n,i}(x)$ during the time period $s,\ldots,s+\ell-1$ implies that there must be at least two distinct clusters during the time period $s,\ldots,s+\ell-1$. Since each cluster ends with an escape, *i.e.*, the occurrence of $Q^{(0)}_{q,n,j}$, then this must have happened at some moment $j \in \{s,\ldots,s+\ell-1\}$ which was then followed by another exceedance at some subsequent instant

$r > j$ where a new cluster is begun. Consequently, we have

$$\mathbb{P}(B^*_{n,s,\ell}(x)) \leq \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}\right)$$

The result follows now at once since

$$\begin{aligned}\left|\mathbb{P}(\mathscr{A}^{s+\ell}_{n,s} \leq x) - \mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x)))\right| &\leq \mathbb{P}\left(\left\{\mathscr{A}^{s+\ell}_{n,s} \leq x\right\} \triangle \mathscr{W}_{s,\ell}(B_{n,i}(x))\right) \\ &= \mathbb{P}(A^*_{n,s,\ell}(x)) + \mathbb{P}(B^*_{n,s,\ell}(x))\end{aligned}$$

If $q = 0$, we start by observing that $\{\mathscr{A}^{s+\ell}_{n,s} \leq x\} \subset \mathscr{W}_{s,\ell}(B_{n,i}(x))$. Then, we note that

$$\left\{\mathscr{A}^{s+\ell}_{n,s} > x\right\} \cap \mathscr{W}_{s,\ell}(B_{n,i}(x)) \subset \bigcup_{j=s}^{s+\ell-1} \bigcup_{r=j+1}^{s+\ell-1} \{X_j > u_{n,j}\} \cap \{X_r > u_{n,r}\}.$$

This is because no entrance in $B_{n,i}(x)$ for $i \in \{s, \ldots, s+\ell-1\}$ implies that there must be at least two exceedances during the time period $s, \ldots, s+\ell-1$.

Consequently, we have

$$\begin{aligned}\left|\mathbb{P}(\mathscr{A}^{s+\ell}_{n,s} \leq x) - \mathbb{P}\left(\mathscr{W}_{s,\ell}(B_{n,i}(x))\right)\right| &= \mathbb{P}\left(\left\{\mathscr{A}^{s+\ell}_{n,s} > x\right\} \cap \mathscr{W}_{s,\ell}(B_{n,i}(x))\right) \\ &\leq \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}(X_j > u_{n,j}, X_r > u_{n,r})\end{aligned}$$

□

As a consequence we obtain an approximation for the Laplace transform of $\mathscr{A}^{s+\ell}_{n,s}$.

Corollary 3.A. *For any $s, \ell \in \mathbb{N}$, $y \geq 0$ and n sufficiently large we have*

$$\begin{aligned}&\left|\mathbb{E}\left(e^{-ya_n\mathscr{A}^{s+\ell}_{n,s}}\right) - \left(1 - \sum_{j=s}^{s+\ell-1} \int_0^\infty y e^{-yx} \mathbb{P}(A_{n,j}(x/a_{n,j}))dx\right)\right| \\ &\qquad \leq 2 \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}\right) + \int_0^\infty y e^{-yx} \delta_{n,s,\ell}(x/a_n)dx\end{aligned}$$

Proof. Using Lemmas 3.2-3.4, for every $x > 0$ we have when $q > 0$

$$\begin{aligned}
&\left|\mathbb{P}\big(\mathscr{A}_{n,s}^{s+\ell} \le x\big) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right)\right| \le \left|\mathbb{P}\big(\mathscr{A}_{n,s}^{s+\ell} \le x\big) - \mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x)))\right| \\
&\quad + |\mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x)) - \mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x)))| + \left|\mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x))) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right)\right| \\
&\le \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \sum_{\kappa=1}^{\lfloor \ell/q \rfloor} \sum_{i=s+\ell-\kappa q}^{s+\ell-1} \mathbb{P}\left(R_{n,i}^{(\kappa)}(x)\right) + \sum_{i=s}^{s+\ell-1} \sum_{\kappa > \lfloor \ell/q \rfloor} \mathbb{P}\left(R_{n,i}^{(\kappa)}(x)\right) \\
&\quad + \sum_{i=1}^{q} \mathbb{P}(B_{n,s+\ell-i}(x)) + \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) \\
&= 2 \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \delta_{n,s,\ell}(x)
\end{aligned}$$

When $q = 0$, we have

$$\begin{aligned}
&\left|\mathbb{P}\big(\mathscr{A}_{n,s}^{s+\ell} \le x\big) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right)\right| \le \left|\mathbb{P}\big(\mathscr{A}_{n,s}^{s+\ell} \le x\big) - \mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x)))\right| \\
&\quad + |\mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x)) - \mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x)))| + \left|\mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x))) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right)\right| \\
&\le \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}(X_j > u_{n,j}, X_r > u_{n,r}) + \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{0,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) \\
&= 2 \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{0,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \delta_{n,s,\ell}(x)
\end{aligned}$$

Since $\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} < 0) = 0$, using integration by parts we have

$$\begin{aligned}
&\mathbb{E}\left(\mathrm{e}^{-y a_n \mathscr{A}_{n,s}^{s+\ell}}\right) = \mathrm{e}^{-y.0}\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} = 0) + \int_0^\infty \mathrm{e}^{-yx} d\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x/a_n) \\
&= \mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} = 0) + \lim_{x\to\infty} \left[\mathrm{e}^{-yx}\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x/a_n) - \mathrm{e}^{-y.0}\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le 0)\right] - \int_0^\infty \mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x/a_n) d\mathrm{e}^{-yx} \\
&= \mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} = 0) - \mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le 0) - \int_0^\infty (-y\mathrm{e}^{-yx})\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x/a_n) dx \\
&= \int_0^\infty y\mathrm{e}^{-yx}\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x/a_n) dx
\end{aligned}$$

Then, using the assumption that $\mathbb{P}(A_{n,j}(x/a_{n,j})) = \mathbb{P}(A_{n,j}(x/a_n))$,

$$\begin{aligned}
&\left|\mathbb{E}\left(\mathrm{e}^{-ya_n\mathscr{A}_{n,s}^{s+\ell}}\right) - \left(1 - \sum_{j=s}^{s+\ell-1}\int_0^\infty y\mathrm{e}^{-yx}\mathbb{P}(A_{n,j}(x/a_{n,j}))dx\right)\right| \\
&= \left|\mathbb{E}\left(\mathrm{e}^{-ya_n\mathscr{A}_{n,s}^{s+\ell}}\right) - \left(1 - \sum_{j=s}^{s+\ell-1}\int_0^\infty y\mathrm{e}^{-yx}\mathbb{P}(A_{n,j}(x/a_n))dx\right)\right| \\
&= \left|\int_0^\infty y\mathrm{e}^{-yx}\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x/a_n)dx - \int_0^\infty y\mathrm{e}^{-yx}\left(1 - \sum_{j=s}^{s+\ell-1}\int_0^\infty y\mathrm{e}^{-yx}\mathbb{P}(A_{n,j}(x/a_n))\right)dx\right| \\
&\le \int_0^\infty y\mathrm{e}^{-yx}\left[2\sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1}\mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \delta_{n,s,\ell}(x/a_n)\right]dx \\
&= 2\sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1}\mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \int_0^\infty y\mathrm{e}^{-yx}\delta_{n,s,\ell}(x/a_n)dx
\end{aligned}$$

□

Next result gives the main induction tool to build the proof of Theorem 2.A.

Lemma 3.5. *Let $s, \ell, t, \varsigma \in \mathbb{N}$ and consider $x_1 \in \mathbb{R}_0^+$, $\underline{x} = (x_2, \ldots, x_\varsigma) \in (\mathbb{R}_0^+)^{\varsigma-1}$, $s + \ell - 1 + t < a_2 < b_2 < a_3 < \ldots < b_{\varsigma-1} < a_\varsigma < b_\varsigma \in \mathbb{N}_0$. For n sufficiently large we have*

$$\begin{aligned}
&\left|\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x_1, \mathscr{A}_{n,a_2}^{b_2} \le x_2, \ldots, \mathscr{A}_{n,a_\varsigma}^{b_\varsigma} \le x_\varsigma) - \mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \le x_1)\mathbb{P}(\mathscr{A}_{n,a_2}^{b_2} \le x_2, \ldots, \mathscr{A}_{n,a_\varsigma}^{b_\varsigma} \le x_\varsigma)\right| \\
&\qquad \le \ell\iota(n,t) + 4\sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1}\mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + 2\delta_{n,s,\ell}(x_1)
\end{aligned}$$

where

$$\iota(n,t) = \sup_{s,\ell\in\mathbb{N}}\max_{i=s,\ldots,s+\ell-1}\left\{\left|\mathbb{P}(A_{n,i}(x_1))\mathbb{P}\big(\cap_{j=2}^{\varsigma}\{\mathscr{A}_{n,a_j}^{b_j} \le x_j\}\big) - \mathbb{P}\big(\cap_{j=2}^{\varsigma}\{\mathscr{A}_{n,a_j}^{b_j} \le x_j\} \cap A_{n,i}(x_1)\big)\right|\right\}. \tag{3.4}$$

Proof. Let

$$\begin{aligned}
A_{x_1,\underline{x}} &:= \{\mathscr{A}_{n,s}^{s+\ell} \le x_1, \mathscr{A}_{n,a_2}^{b_2} \le x_2, \ldots, \mathscr{A}_{n,a_\varsigma}^{b_\varsigma} \le x_\varsigma\}, \qquad B_{x_1} := \{\mathscr{A}_{n,s}^{s+\ell} \le x_1\} \\
\tilde{A}_{x_1,\underline{x}} &:= \mathscr{W}_{s,\ell}(A_{n,i}(x_1)) \cap \{\mathscr{A}_{n,a_2}^{b_2} \le x_2, \ldots, \mathscr{A}_{n,a_\varsigma}^{b_\varsigma} \le x_\varsigma\}, \quad \tilde{B}_{x_1} := \mathscr{W}_{s,\ell}(A_{n,i}(x_1)), \\
D^{\underline{x}} &:= \{\mathscr{A}_{n,a_2}^{b_2} \le x_2, \ldots, \mathscr{A}_{n,a_\varsigma}^{b_\varsigma} \le x_\varsigma\}.
\end{aligned}$$

If $x_1 > 0$, by Lemmas 3.2 and 3.4 we have

$$\begin{aligned}
\left|\mathbb{P}(B_{x_1}) - \mathbb{P}(\tilde{B}_{x_1})\right| &\leq \left|\mathbb{P}(\mathscr{A}_{n,s}^{s+\ell} \leq x_1) - \mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x_1)))\right| + |\mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x_1))) - \mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x_1)))| \\
&\leq \left|\mathbb{P}(\{\mathscr{A}_{n,s}^{s+\ell} \leq x_1\} \triangle \mathscr{W}_{s,\ell}(B_{n,i}(x_1)))\right| + |\mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x_1)) \setminus \mathscr{W}_{s,\ell}(B_{n,i}(x_1)))| \\
&\leq \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \sum_{\kappa=1}^{\lfloor \ell/q \rfloor} \sum_{i=s+\ell-\kappa q}^{s+\ell-1} \mathbb{P}\left(R_{n,i}^{(\kappa)}(x_1)\right) \\
&\quad + \sum_{i=s}^{s+\ell-1} \sum_{\kappa > \lfloor \ell/q \rfloor} \mathbb{P}\left(R_{n,i}^{(\kappa)}(x_1)\right) + \sum_{i=1}^{q} \mathbb{P}(B_{n,s+\ell-i}(x_1)) \\
&= \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \delta_{n,s,\ell}(x_1) \qquad (3.5)
\end{aligned}$$

and also

$$\begin{aligned}
\left|\mathbb{P}(A_{x_1}) - \mathbb{P}(\tilde{A}_{x_1})\right| &\leq \left|\mathbb{P}(\{\mathscr{A}_{n,s}^{s+\ell} \leq x_1\} \cap D^{\underline{x}}) - \mathbb{P}(\mathscr{W}_{s,\ell}(B_{n,i}(x_1)) \cap D^{\underline{x}})\right| + |\mathbb{P}\left((\mathscr{W}_{s,\ell}(A_{n,i}(x_1)) \setminus \mathscr{W}_{s,\ell}(B_{n,i}(x_1))) \cap D^{\underline{x}}\right)| \\
&\leq \left|\mathbb{P}\left((\{\mathscr{A}_{n,s}^{s+\ell} \leq x_1\} \triangle \mathscr{W}_{s,\ell}(B_{n,i}(x_1)) \cap D^{\underline{x}}\right)\right| + |\mathbb{P}\left((\mathscr{W}_{s,\ell}(A_{n,i}(x_1)) \setminus \mathscr{W}_{s,\ell}(B_{n,i}(x_1))) \cap D^{\underline{x}}\right)| \\
&\leq \left|\mathbb{P}(\{\mathscr{A}_{n,s}^{s+\ell} \leq x_1\} \triangle \mathscr{W}_{s,\ell}(B_{n,i}(x_1))\right| + |\mathbb{P}(\mathscr{W}_{s,\ell}(A_{n,i}(x_1)) \setminus \mathscr{W}_{s,\ell}(B_{n,i}(x_1)))| \\
&\leq \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \delta_{n,s,\ell}(x_1) \qquad (3.6)
\end{aligned}$$

If $x_1 = 0$, we notice that $\{\mathscr{A}_{n,s}^{s+\ell} \leq x_1\} = \{\mathscr{A}_{n,s}^{s+\ell} = 0\} = \{X_s \leq u_{n,s}, \ldots, X_{s+\ell-1} \leq u_{n,s+\ell-1}\} = \mathscr{W}_{s,\ell}(B_{n,i}(0))$, so estimates (3.5) and (3.6) are still valid by Lemma 3.2.

Adapting the proof of Lemma 3.3, it follows that

$\left|\mathbb{P}(\tilde{A}_{x_1,\underline{x}}) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right) \mathbb{P}(D^{\underline{x}})\right| \leq Err$, where

$$Err = \left|\sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\mathbb{P}(D^{\underline{x}}) - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x) \cap D^{\underline{x}})\right| + \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right)$$

Now, since, by definition of $\iota(n,t)$,

$$\begin{aligned}
&\left|\sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\mathbb{P}(D^{\underline{x}}) - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x) \cap D^{\underline{x}})\right| \\
&\qquad \leq \sum_{i=s}^{s+\ell-1} |\mathbb{P}(A_{n,i}(x))\mathbb{P}(D^{\underline{x}}) - \mathbb{P}(A_{n,i}(x) \cap D^{\underline{x}})| \leq \ell\iota(n,t),
\end{aligned}$$

we conclude that

$$\left|\mathbb{P}(\tilde{A}_{x_1,\underline{x}}) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x))\right) \mathbb{P}(D^{\underline{x}})\right| \leq \ell\iota(n,t) + \sum_{j=s}^{s+\ell-1} \sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) \qquad (3.7)$$

Also, by Lemma 3.3 we have

$$\left|\mathbb{P}(\tilde{B}_{x_1})\mathbb{P}(D^{\underline{x}}) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x_1))\right)\mathbb{P}(D^{\underline{x}})\right| \leq \sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}\right) \quad (3.8)$$

Putting together the estimates (3.5)-(3.8) we get

$$\begin{aligned}|\mathbb{P}(A_{x_1,\underline{x}}) - \mathbb{P}(B_{x_1})\mathbb{P}(D^{\underline{x}})| &\leq \left|\mathbb{P}(A_{x_1,\underline{x}}) - \mathbb{P}(\tilde{A}_{x_1,\underline{x}})\right| + \left|\mathbb{P}(\tilde{A}_{x_1,\underline{x}}) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x_1))\right)\mathbb{P}(D^{\underline{x}})\right| \\ &+ \left|\mathbb{P}(\tilde{B}_{x_1})\mathbb{P}(D^{\underline{x}}) - \left(1 - \sum_{i=s}^{s+\ell-1} \mathbb{P}(A_{n,i}(x_1))\right)\mathbb{P}(D^{\underline{x}})\right| + \left|\mathbb{P}(B_{x_1}) - \mathbb{P}(\tilde{B}_{x_1})\right|\mathbb{P}(D^{\underline{x}}) \\ &\leq \ell\iota(n,t) + 4\sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1} \mathbb{P}\left(Q^{(0)}_{q,n,j} \cap \{X_r > u_{n,r}\}\right) + 2\delta_{n,s,\ell}(x_1)\end{aligned}$$

□

Let us consider a function $F : (\mathbb{R}_0^+)^n \to \mathbb{R}$ which is continuous on the right in each variable separately and such that for each $R = (a_1, b_1] \times \ldots \times (a_n, b_n] \subset (\mathbb{R}_0^+)^n$ we have

$$\mu_F(R) := \sum_{c_i \in \{a_i, b_i\}} (-1)^{\#\{i \in \{1,\ldots,n\} : c_i = a_i\}} F(c_1, \ldots, c_n) \geq 0$$

Such F is called an *n-dimensional Stieltjes measure function* and such μ_F has a unique extension to the Borel σ-algebra in $(\mathbb{R}_0^+)^n$, which is called the *Lebesgue-Stieltjes measure* associated to F.

For each $I \subset \{1, \ldots, n\}$, let $F_I(\underline{x}) := F(\delta_1 x_1, \ldots, \delta_n x_n)$, where $\delta_i = \begin{cases} 1 & \text{if } i \in I \\ 0 & \text{if } i \notin I \end{cases}$

If F is an n-dimensional Stieltjes measure function, it is easy to see that F_I is also an n-dimensional Stieltjes measure function, which has an associated Lebesgue-Stieltjes measure μ_{F_I}. We will use the following proposition, proved in [FFMa18, Section 4]:

Proposition 3.B. *Given $n \in \mathbb{N}$, $I \subset \{1, \ldots, n\}$ and two functions $F, G : (\mathbb{R}_0^+)^n \to \mathbb{R}$ such that F is a bounded n-dimensional Stieltjes measure function, let*

$$\int G(\underline{x}) dF_I(\underline{x}) := \begin{cases} G(0,\ldots,0)F(0,\ldots,0) & \text{for } I = \emptyset \\ \int G(\underline{x}) d\mu_{F_I} & \text{for } I \neq \emptyset \end{cases}$$

where μ_{F_I} is the Lebesgue-Stieltjes measure associated to F_I. Then,

$$\int_0^\infty \cdots \int_0^\infty e^{-y_1x_1 - \ldots - y_nx_n} F(\underline{x}) dx_1 \ldots dx_n = \frac{1}{y_1 \ldots y_n} \sum_{I \subset \{1,\ldots,n\}} \int e^{-\sum_{i \in I} y_i x_i} dF_I(\underline{x})$$

Corollary 3.C. *Let $s, \ell, t, \varsigma \in \mathbb{N}$ and consider $y_1, y_2, \ldots, y_\varsigma \in \mathbb{R}_0^+$, $s + \ell - 1 + t < a_2 < b_2 < a_3 < \ldots < b_{\varsigma-1} < a_\varsigma < b_\varsigma \in \mathbb{N}_0$. For n sufficiently large we have*

$$\mathbb{E}\left(e^{-y_1 a_n \mathscr{A}^{s+\ell}_{n,s} - y_2 a_n \mathscr{A}^{b_2}_{n,a_2} - \ldots - y_\varsigma a_n \mathscr{A}^{b_\varsigma}_{n,a_\varsigma}}\right) = \mathbb{E}\left(e^{-y_1 a_n \mathscr{A}^{s+\ell}_{n,s}}\right)\mathbb{E}\left(e^{-y_2 a_n \mathscr{A}^{b_2}_{n,a_2} - \ldots - y_\varsigma a_n \mathscr{A}^{b_\varsigma}_{n,a_\varsigma}}\right) + Err$$

where

$$|Err| \le \ell\iota(n,t) + 4\sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1}\mathbb{P}\left(Q^{(0)}_{q,n,j}\cap\{X_r > u_{n,r}\}\right) + 2\int_0^\infty y_1 e^{-y_1 x}\delta_{n,s,\ell}(x/a_n)dx$$

and $\iota(n,t)$ *is given by* (3.4).

Proof. Using the same notation as in the proof of Lemma 3.5, let $F^{(A)}(x_1,\dots,x_\varsigma) = \mathbb{P}(A_{x_1,\underline{x}})$, $F^{(B)}(x_1) = \mathbb{P}(B_{x_1})$ and $F^{(D)}(x_2,\dots,x_\varsigma) = \mathbb{P}(D^{\underline{x}})$. Then, $F^{(A)}$, $F^{(B)}$ and $F^{(D)}$ are both bounded Stieltjes measure functions, with

$$\mu_{F^{(A)}}(U_1) = \mathbb{P}\left((a_n\mathscr{A}^{s+\ell}_{n,s}, a_n\mathscr{A}^{b_2}_{n,a_2},\dots,a_n\mathscr{A}^{b_\varsigma}_{n,a_\varsigma})\in U_1\right)$$

$$\mu_{F^{(B)}}(U_2) = \mathbb{P}(a_n\mathscr{A}^{s+\ell}_{n,s}\in U_2) \qquad \mu_{F^{(D)}}(U_3) = \mathbb{P}\left((a_n\mathscr{A}^{b_2}_{n,a_2},\dots,a_n\mathscr{A}^{b_\varsigma}_{n,a_\varsigma})\in U_3\right)$$

where U_1, U_2 and U_3 are Borel sets in $(\mathbb{R}^+_0)^\varsigma$, $\mathbb{R}^+_0$ and $(\mathbb{R}^+_0)^{\varsigma-1}$, respectively.

Therefore, we can apply the previous proposition and we obtain

$$\begin{aligned}
&\mathbb{E}\left(\mathrm{e}^{-y_1a_n\mathscr{A}^{s+\ell}_{n,s}-y_2a_n\mathscr{A}^{b_2}_{n,a_2}-\dots-y_\varsigma a_n\mathscr{A}^{b_\varsigma}_{n,a_\varsigma}}\right) - \mathbb{E}\left(\mathrm{e}^{-y_1a_n\mathscr{A}^{s+\ell}_{n,s}}\right)\mathbb{E}\left(\mathrm{e}^{-y_2a_n\mathscr{A}^{b_2}_{n,a_2}-\dots-y_\varsigma a_n\mathscr{A}^{b_\varsigma}_{n,a_\varsigma}}\right)\\
&= \sum_{I\subset\{1,\dots,\varsigma\}}\int \mathrm{e}^{-\sum_{i\in I}y_ia_nx_i}d(F^{(A)})_I(x_1,\dots,x_\varsigma)\\
&- \sum_{I\subset\{1\}}\int \mathrm{e}^{-\sum_{i\in I}y_ia_nx_i}d(F^{(B)})_I(x_1)\sum_{I\subset\{2,\dots,\varsigma\}}\int \mathrm{e}^{-\sum_{i\in I}y_ia_nx_i}d(F^{(D)})_I(x_2,\dots,x_\varsigma)\\
&= y_1\dots y_\varsigma a_n^\varsigma\int_0^\infty\dots\int_0^\infty \mathrm{e}^{-y_1a_nx_1-\dots-y_\varsigma a_nx_\varsigma}F^{(A)}(x_1,\dots,x_\varsigma)dx_1\dots dx_\varsigma\\
&- \left(y_1a_n\int_0^\infty \mathrm{e}^{-y_1a_nx_1}F^{(B)}(x_1)dx_1\right)\left(y_2\dots y_\varsigma a_n^{\varsigma-1}\int_0^\infty\dots\int_0^\infty \mathrm{e}^{-y_2a_nx_2-\dots-y_\varsigma a_nx_\varsigma}F^{(D)}(x_2,\dots,x_\varsigma)dx_2\dots dx_\varsigma\right)\\
&= y_1\dots y_\varsigma a_n^\varsigma\int_0^\infty\dots\int_0^\infty \mathrm{e}^{-y_1a_nx_1-\dots-y_\varsigma a_nx_\varsigma}(F^{(A)}-F^{(B)}F^{(D)})(x_1,\dots,x_\varsigma)dx_1\dots dx_\varsigma
\end{aligned}$$

Hence, using Lemma 3.5,

$$\begin{aligned}
&\left|\mathbb{E}\left(\mathrm{e}^{-y_1a_n\mathscr{A}^{s}_{n,0}-y_2a_n\mathscr{A}^{b_2}_{n,a_2}-\dots-y_\varsigma a_n\mathscr{A}^{b_\varsigma}_{n,a_\varsigma}}\right) - \mathbb{E}\left(\mathrm{e}^{-y_1a_n\mathscr{A}^{s}_{n,0}}\right)\mathbb{E}\left(\mathrm{e}^{-y_2a_n\mathscr{A}^{b_2}_{n,a_2}-\dots-y_\varsigma a_n\mathscr{A}^{b_\varsigma}_{n,a_\varsigma}}\right)\right|\\
&\qquad\le y_1\dots y_\varsigma a_n^\varsigma\int_0^\infty\dots\int_0^\infty \mathrm{e}^{-y_1ax_1-\dots-y_\varsigma ax_\varsigma}\left|\mathbb{P}(A_{x_1,\underline{x}})-\mathbb{P}(B_{x_1})\mathbb{P}(D^{\underline{x}})\right|dx_1\dots dx_\varsigma\\
&\qquad\le \ell\iota(n,t) + 4\sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1}\mathbb{P}\left(Q^{(0)}_{q,n,j}\cap\{X_r > u_{n,r}\}\right) + 2y_1a_n\int_0^\infty \mathrm{e}^{-y_1a_nx_1}\delta_{n,s,\ell}(x_1)dx_1\\
&\qquad= \ell\iota(n,t) + 4\sum_{j=s}^{s+\ell-1}\sum_{r=j+1}^{s+\ell-1}\mathbb{P}\left(Q^{(0)}_{q,n,j}\cap\{X_r > u_{n,r}\}\right) + 2\int_0^\infty y_1\mathrm{e}^{-y_1x}\delta_{n,s,\ell}(x/a_n)dx
\end{aligned}$$

□

Proposition 3.D. *Let* $X_0, X_1,\dots$ *be given by* (2.1)*, let* $J\in\mathcal{R}$ *be such that* $J = \bigcup_{\ell=1}^\varsigma I_\ell$ *where* $I_j = [a_j,b_j)\in\mathcal{S}$*,* $j = 1,\dots,\varsigma$ *and* $a_1 < b_1 < a_2 < \dots < b_{\varsigma-1} < a_\varsigma < b_\varsigma$*, let* $u_{n,i}$ *be real-valued boundaries satisfying* (2.2) *and* (2.3)*, let* $H := \lceil\sup\{x : x\in J\}\rceil = \lceil b_\varsigma\rceil$ *and let* $(a_n)_{n\in\mathbb{N}}$ *be a normalising sequence,* $a_{n,j}$ *normalising factors and* π *a probability distribution as in* (2.A).

Assume that $Д_q(u_{n,i})^*$, $Д'_q(u_{n,i})^*$ *and* $ULC_q(u_{n,i})^*$ *hold, for some* $q \in \mathbb{N}_0$. *Consider the partition of* $[0, Hn)$ *into blocks of length* $\ell_{H,n,j}$, $J_1 = [\mathcal{L}_{H,n,0}, \mathcal{L}_{H,n,1})$, $J_2 = [\mathcal{L}_{H,n,1}, \mathcal{L}_{H,n,2})$, ..., $J_{k_n} = [\mathcal{L}_{H,n,k_n-1}, \mathcal{L}_{H,n,k_n})$, $J_{k_n+1} = [\mathcal{L}_{H,n,k_n}, Hn)$. *Let* n *be sufficiently large so that* $L_{H,n} := \max\{\ell_{H,n,j}, j = 1, \dots, k_n\} < \frac{n}{2} \inf_{j\in\{1,\dots,\varsigma\}}\{b_j - a_j\}$ *and, finally, let* $\mathscr{S}_\ell$ *be the number of blocks* $J_i, i > 1$ *contained in* nI_ℓ, *that is,*

$$\mathscr{S}_\ell := \#\{i \in \{2, \dots, k_n\} : J_i \subset nI_\ell\}.$$

Note that, by definition of $L_{H,n}$, *we must have* $\mathscr{S}_\ell > 1$ *for every* $\ell \in \{1, \dots, \varsigma\}$.

Then, for all $y_1, y_2, \dots, y_\varsigma \in \mathbb{R}_0^+$, *we have*

$$\mathbb{E}\left(e^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n(nI_\ell)}\right) - \prod_{\ell=1}^{\varsigma} \prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\left(e^{-y_\ell a_n \mathscr{A}_n(J_i)}\right) \xrightarrow[n\to\infty]{} 0$$

Proof. Without loss of generality, we can assume that $y_1, y_2, \dots, y_\varsigma \in \mathbb{R}^+$, because if we had $y_j = 0$ for some $j = 1, \dots, \varsigma$ then we could consider $J = \bigcup_{\ell=1}^{j-1} I_\ell \cup \bigcup_{\ell=j+1}^{\varsigma} I_\ell$ instead. Also, we can assume that $a_1 > 0$. Let $\hat{y} := \inf\{y_j : j = 1, \dots, \varsigma\} > 0$ and $\hat{Y} := \sup\{y_j : j = 1, \dots, \varsigma\}$. We cut each J_i into two blocks:

$$J_i^* := [\mathcal{L}_{H,n,i-1}, \mathcal{L}_{H,n,i} - t_{H,n,i}) \text{ and } J_i' := J_i \setminus J_i^*$$

Note that $|J_i^*| = \ell_{H,n,i} - t_{H,n,i}$ and $|J_i'| = t_{H,n,i}$.

For each $\ell \in \{1, \dots, \varsigma\}$, we define $i_\ell := \min\{i \in \{2, \dots, k_n\} : J_i \subset nI_\ell\}$. Hence, it follows that $J_{i_\ell}, J_{i_\ell+1}, \dots, J_{i_\ell+\mathscr{S}_\ell-1} \subset nI_\ell$ and

$$\mathcal{L}_{H,n,i_\ell+\mathscr{S}_\ell-1} - \mathcal{L}_{H,n,i_\ell-1} = \sum_{j=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \ell_{H,n,j} \sim n|I_\ell| \tag{3.9}$$

First of all, recall that for every $0 \le x_i, z_i \le 1$, we have

$$\left|\prod x_i - \prod z_i\right| \le \sum |x_i - z_i|. \tag{3.10}$$

We start by making the following approximation, in which we use (3.10),

$$\begin{aligned}\left|\mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n(nI_\ell)}\right) - \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell \sum_{j=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} a_n \mathscr{A}_n(J_j)}\right)\right| &\le \mathbb{E}\left(1 - \mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n\left(nI_\ell \setminus \cup_{j=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} J_j\right)}\right)\\ &\le \mathbb{E}\left(1 - \mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n(J_{i_\ell-1} \cup J_{i_\ell+\mathscr{S}_\ell})}\right)\\ &\le \varsigma K \mathbb{E}\left(1 - \mathrm{e}^{-a_n \mathscr{A}_n(J_{i_\ell-1})}\right) + \varsigma K \mathbb{E}\left(1 - \mathrm{e}^{-a_n \mathscr{A}_n(J_{i_\ell+\mathscr{S}_\ell})}\right),\end{aligned}$$

where $\max\{y_1, \dots, y_\varsigma\} \le K \in \mathbb{N}$. In order to show that we are allowed to use the above approximation we just need to check that $\mathbb{E}\left(1 - \mathrm{e}^{-a_n \mathscr{A}_n(J_i)}\right) \to 0$ as $n \to \infty$ for every $i = 1, \dots, k_n + 1$. By Corollary 3.A we have for $i = 1, \dots, k_n$

$$\mathbb{E}\left(\mathrm{e}^{-a_n \mathscr{A}_n(J_i)}\right) = \mathbb{E}\left(\mathrm{e}^{-a_n \mathscr{A}_{n,\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}}}\right) = 1 - \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \int_0^\infty \mathrm{e}^{-x} \mathbb{P}(A_{n,j}(x/a_{n,j}))dx + Err, \tag{3.11}$$

where

$$|Err| \le 2 \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \sum_{r=j+1}^{\mathcal{L}_{H,n,i}-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \int_0^\infty \mathrm{e}^{-x} \delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}}(x/a_n)dx \to 0$$

as $n \to \infty$ by $Д_q'(u_{n,i})^*$ and $ULC_q(u_{n,i})$. Since

$$\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \int_0^\infty \mathrm{e}^{-x}\mathbb{P}(A_{n,j}(x/a_{n,j}))dx \le \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \int_0^\infty \mathrm{e}^{-x}\mathbb{P}(X_j > u_{n,j})dx = \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \bar{F}(u_{n,j}) \le \frac{F_{H,n}^*}{k_n}$$

we get $\mathbb{E}\left(\mathrm{e}^{-a_n\mathscr{A}_n(J_i)}\right) \xrightarrow[n\to\infty]{} 1$ by (2.3).

If $i = k_n + 1$ then

$$\mathbb{E}\left(\mathrm{e}^{-a_n\mathscr{A}_n(J_i)}\right) = \mathbb{E}\left(\mathrm{e}^{-a_n\mathscr{A}_{n,\mathcal{L}_{H,n,k_n}}^{Hn}}\right) = 1 - \sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1} \int_0^\infty \mathrm{e}^{-x}\mathbb{P}(A_{n,j}(x/a_n))dx + Err, \quad (3.12)$$

where

$$|Err| \le 2 \sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1} \sum_{r=j+1}^{Hn-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \int_0^\infty \mathrm{e}^{-x} \delta_{n,\mathcal{L}_{H,n,k_n},Hn-\mathcal{L}_{H,n,k_n}}(x/a_n)dx \to 0$$

as $n \to \infty$ by $Д_q'(u_{n,i})^*$ and $ULC_q(u_{n,i})$. Since, by (3.2),

$$\sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1} \int_0^\infty \mathrm{e}^{-x}\mathbb{P}(A_{n,j}(x/a_{n,j}))dx \le \sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1} \int_0^\infty \mathrm{e}^{-x}\mathbb{P}(X_j > u_{n,j})dx = \sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1} \bar{F}(u_{n,j}) \le k_n \bar{F}_{n,\max}(H)$$

we get $\mathbb{E}\left(\mathrm{e}^{-a_n\mathscr{A}_n(J_{k_n+1})}\right) \xrightarrow[n\to\infty]{} 1$ by (2.9).

Now, we proceed with another approximation which consists of replacing J_i by J_i^*. Using (3.10) we have

$$\begin{aligned}
\left|\mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell \sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} a_n\mathscr{A}_n(J_i)}\right) - \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell \sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} a_n\mathscr{A}_n(J_i^*)}\right)\right| &\le \mathbb{E}\left(1 - \mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell \sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} a_n\mathscr{A}_n(J_i')}\right) \\
&\le K \sum_{\ell=1}^{\varsigma} \mathbb{E}\left(1 - \mathrm{e}^{-\sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} a_n\mathscr{A}_n(J_i')}\right) \le K \sum_{\ell=1}^{\varsigma} \sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\left(1 - \mathrm{e}^{-a_n\mathscr{A}_n(J_i')}\right) \\
&\le K \sum_{i=1}^{k_n} \mathbb{E}\left(1 - \mathrm{e}^{-a_n\mathscr{A}_n(J_i')}\right)
\end{aligned}$$

Now, we must show that $\sum_{i=1}^{k_n} \mathbb{E}\left(1 - \mathrm{e}^{-a_n\mathscr{A}_n(J_i')}\right) \to 0$, as $n \to \infty$, in order for the approximation to make sense. By Corollary 3.A we have

$$\mathbb{E}\left(\mathrm{e}^{-a_n\mathscr{A}_n(J_i')}\right) = \mathbb{E}\left(\mathrm{e}^{-a_n\mathscr{A}_{n,\mathcal{L}_{H,n,i}-t_{H,n,i}}^{\mathcal{L}_{H,n,i}}}\right) = 1 - \sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}}^{\mathcal{L}_{H,n,i}-1} \int_0^\infty \mathrm{e}^{-x}\mathbb{P}(A_{n,j}(x/a_{n,j}))dx + Err, \quad (3.13)$$

where

$$\sum_{i=1}^{k_n}|Err| \leq 2\sum_{i=1}^{k_n}\sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}}^{\mathcal{L}_{H,n,i}-1}\sum_{r=j+1}^{\mathcal{L}_{H,n,i}-1}\mathbb{P}\left(Q_{q,n,j}^{(0)}\cap\{X_r>u_{n,r}\}\right)$$
$$+\sum_{i=1}^{k_n}\int_0^\infty \mathrm{e}^{-x}\delta_{\mathcal{L}_{H,n,i}-t_{H,n,i},t_{H,n,i},n}(x/a_n)dx \to 0$$

as $n\to\infty$ by $Д_q'(u_{n,i})^*$ and $ULC_q(u_{n,i})$. We get, by (2.9) as well,

$$\sum_{i=1}^{k_n}\mathbb{E}\left(1-\mathrm{e}^{-a_n\mathscr{A}_n(J_i')}\right)\sim\sum_{i=1}^{k_n}\sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}}^{\mathcal{L}_{H,n,i}-1}\int_0^\infty \mathrm{e}^{-x}\mathbb{P}(A_{n,j}(x/a_{n,j}))dx$$
$$\leq\sum_{i=1}^{k_n}\sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})\leq\sum_{i=1}^{k_n}\varepsilon(H,n)\frac{F_{H,n}^*}{k_n}=k_n(t_n^*+1)\bar{F}_{n,\max}(H)\xrightarrow[n\to\infty]{}0$$

Let us fix now some $\hat{\ell}\in\{1,\dots,\varsigma\}$ and $i\in\{i_{\hat{\ell}},\dots,i_{\hat{\ell}}+\mathscr{S}_{\hat{\ell}}-1\}$. Let $M_i=y_{\hat{\ell}}\sum_{j=i}^{i_{\hat{\ell}}+\mathscr{S}_{\hat{\ell}}-1}a_n\mathscr{A}_n(J_j^*)$ and $L_{\hat{\ell}}=\sum_{\ell=\hat{\ell}+1}^{\varsigma}y_\ell\sum_{j=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}a_n\mathscr{A}_n(J_j^*)$. Using Corollary 3.C along with the facts that $\iota(n,t)\leqslant\gamma_i(n,t)$ and $\gamma_i(n,t)$ is decreasing in t, we obtain

$$\left|\mathbb{E}\left(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_{i_{\hat{\ell}}}^*)-M_{i_{\hat{\ell}}+1}-L_{\hat{\ell}}}\right)-\mathbb{E}\left(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_{i_{\hat{\ell}}}^*)}\right)\mathbb{E}\left(\mathrm{e}^{-M_{i_{\hat{\ell}}+1}-L_{\hat{\ell}}}\right)\right|\leq\Upsilon_{n,i_{\hat{\ell}}},$$

where

$$\Upsilon_{n,i}=t_{H,n,i}\gamma_i(n,t_{H,n,i})+4\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i-1}+t_{H,n,i}-1}\sum_{r=j+1}^{\mathcal{L}_{H,n,i-1}+t_{H,n,i}-1}\mathbb{P}\left(Q_{q,n,j}^{(0)}\cap\{X_r>u_{n,r}\}\right)$$
$$+2\int_0^\infty y_{\hat{\ell}}\mathrm{e}^{-y_{\hat{\ell}}x}\delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}-t_{H,n,i}}(x/a_n)dx$$
$$\leq\ell_{H,n,i}\gamma_i(n,t_n^*)+4\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\sum_{r=j+1}^{\mathcal{L}_{H,n,i}-1}\mathbb{P}\left(Q_{q,n,j}^{(0)}\cap\{X_r>u_{n,r}\}\right)$$
$$+2\hat{Y}\int_0^\infty \mathrm{e}^{-\hat{y}x}\delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}-t_{H,n,i}}(x/a_n)dx$$

Since $\mathbb{E}\left(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_i^*)}\right)\leq 1$ for any $i\in\{1,\dots,k_n\}$, it follows by the same argument that

$$\left|\mathbb{E}\Big(\mathrm{e}^{-M_{i_{\hat{\ell}}}-L_{\hat{\ell}}}\Big)-\mathbb{E}\Big(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_{i_{\hat{\ell}}}^*)}\Big)\mathbb{E}\Big(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_{i_{\hat{\ell}}+1}^*)}\Big)\mathbb{E}\Big(\mathrm{e}^{-M_{i_{\hat{\ell}}+2}-L_{\hat{\ell}}}\Big)\right|$$
$$\leq\left|\mathbb{E}\Big(\mathrm{e}^{-M_{i_{\hat{\ell}}}-L_{\hat{\ell}}}\Big)-\mathbb{E}\Big(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_{i_{\hat{\ell}}}^*)}\Big)\mathbb{E}\Big(\mathrm{e}^{-M_{i_{\hat{\ell}}+1}-L_{\hat{\ell}}}\Big)\right|$$
$$+\mathbb{E}\Big(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_{i_{\hat{\ell}}}^*)}\Big)\left|\mathbb{E}\Big(\mathrm{e}^{-M_{i_{\hat{\ell}}+1}-L_{\hat{\ell}}}\Big)-\mathbb{E}\Big(\mathrm{e}^{-y_{\hat{\ell}}a_n\mathscr{A}_n(J_{i_{\hat{\ell}}+1}^*)}\Big)\mathbb{E}\Big(\mathrm{e}^{-M_{i_{\hat{\ell}}+2}-L_{\hat{\ell}}}\Big)\right|$$
$$\leq\Upsilon_{n,i_{\hat{\ell}}}+\Upsilon_{n,i_{\hat{\ell}}+1}$$

Hence, proceeding inductively with respect to $i \in \{i_{\hat{\ell}}, \dots, i_{\hat{\ell}} + \mathscr{S}_{\hat{\ell}} - 1\}$, we obtain

$$\left|\mathbb{E}\Big(\mathrm{e}^{-M_{i_{\hat{\ell}}}-L_{\hat{\ell}}}\Big) - \prod_{j=i_{\hat{\ell}}}^{i_{\hat{\ell}}+\mathscr{S}_{\hat{\ell}}-1} \mathbb{E}\Big(\mathrm{e}^{-y_{\hat{\ell}} a_n \mathscr{A}_n(J_j^*)}\Big)\mathbb{E}\Big(\mathrm{e}^{-L_{\hat{\ell}}}\Big)\right| \le \sum_{i=i_{\hat{\ell}}}^{i_{\hat{\ell}}+\mathscr{S}_{\hat{\ell}}-1} \Upsilon_{n,i}$$

In the same way, if we proceed inductively with respect to $\hat{\ell} \in \{1, \dots, \varsigma\}$, we get

$$\left|\mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell \sum_{j=i_\ell}^{i_\ell+\mathscr{S}_\ell} a_n \mathscr{A}_n(J_j^*)}\right) - \prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\Big(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i^*)}\Big)\right| \le \sum_{\ell=1}^{\varsigma}\sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \Upsilon_{n,i}$$

$$\le \sum_{i=1}^{k_n} \Upsilon_{n,i} \le Hn\gamma_i(n,t_n^*) + 4\sum_{i=1}^{k_n}\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\sum_{r=j+1}^{\mathcal{L}_{H,n,i}-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right)$$

$$+ 2\hat{Y}\sum_{i=1}^{k_n}\int_0^\infty \mathrm{e}^{-\hat{y}x}\delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}-t_{H,n,i}}(x/a_n)dx \to 0$$

as $n \to \infty$, by $Д_q(u_{n,i})^*$, $Д_q'(u_{n,i})^*$ and $ULC_q(u_{n,i})$.

Using (3.10) again, we have the final approximation

$$\left|\prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\Big(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i)}\Big) - \prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\Big(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i^*)}\Big)\right| \le K\sum_{\ell=1}^{\varsigma}\sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\left(1 - \mathrm{e}^{-a_n \mathscr{A}_n(J_i')}\right)$$

$$\le K\sum_{i=1}^{k_n} \mathbb{E}\left(1 - \mathrm{e}^{-a_n \mathscr{A}_n(J_i')}\right).$$

We have already proved that $\sum_{i=1}^{k_n} \mathbb{E}\left(1 - \mathrm{e}^{-a_n \mathscr{A}_n(J_i')}\right) \to 0$ as $n \to \infty$, so we only need to gather all the approximations to finally obtain the stated result.

□

Proof of Theorem 2.A. In order to prove convergence of $a_n A_n$ to a process A, it is sufficient to show that for any ς disjoint intervals $I_1, I_2, \dots, I_\varsigma \in \mathcal{S}$, the joint distribution of $a_n A_n$ over these intervals converges to the joint distribution of A over the same intervals, *i.e.*,

$$(a_n A_n(I_1), a_n A_n(I_2), \dots, a_n A_n(I_\varsigma)) \xrightarrow[n\to\infty]{} (A(I_1), A(I_2), \dots, A(I_\varsigma)),$$

which will be the case if the corresponding joint Laplace transforms converge. Hence, we only need to show that

$$\psi_{a_n A_n}(y_1, y_2, \dots, y_\varsigma) \to \psi_A(y_1, y_2, \dots, y_\varsigma) = \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell A(I_\ell)}\right), \quad \text{as } n \to \infty,$$

for every ς non-negative values $y_1, y_2, \dots, y_\varsigma$, each choice of ς disjoint intervals $I_1, I_2, \dots, I_\varsigma \in \mathcal{S}$ and each $\varsigma \in \mathbb{N}$. Note that $\psi_{a_n A_n}(y_1, y_2, \dots, y_\varsigma) = \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n A_n(I_\ell)}\right) = \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n(v_n I_\ell)}\right)$

and

$$\begin{aligned}\left|\mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n(v_n I_\ell)}\right) - \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell A(I_\ell)}\right)\right| &\leq \left|\mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n(v_n I_\ell)}\right) - \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n\left(\frac{n}{\tau} I_\ell\right)}\right)\right| \\ &+ \left|\mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell a_n \mathscr{A}_n\left(\frac{n}{\tau} I_\ell\right)}\right) - \prod_{\ell=1}^{\varsigma} \prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\left(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i)}\right)\right| \\ &+ \left|\prod_{\ell=1}^{\varsigma} \prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\left(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i)}\right) - \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma} y_\ell A(I_\ell)}\right)\right|\end{aligned}$$

where $J_1, J_2, \ldots, J_{k_n+1}$ are the elements of the partition of $[0, Hn)$ given by Proposition 3.D, with $J = \bigcup_{\ell=1}^{\varsigma} \frac{1}{\tau} I_\ell$. Since $v_n \sim \frac{n}{\tau}$, the first term on the right goes to 0 as $n \to \infty$. By Proposition 3.D, the second term on the right also goes to 0 as $n \to \infty$. Finally, by Corollary 3.A, we have

$$\mathbb{E}\left(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i)}\right) = 1 - \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \int_0^\infty y_\ell \mathrm{e}^{-y_\ell x} \mathbb{P}(A_{n,j}(x/a_{n,j}))dx + Err$$

where

$$|Err| \leq 2 \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \sum_{r=j+1}^{\mathcal{L}_{H,n,i}-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \int_0^\infty y_\ell \mathrm{e}^{-y_\ell x} \delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}}(x/a_n)dx$$

Using (3.10), we have that

$$\begin{aligned}&\left|\prod_{\ell=1}^{\varsigma} \prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \mathbb{E}\left(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i)}\right) - \prod_{\ell=1}^{\varsigma} \prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \left(1 - \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \int_0^\infty y_\ell \mathrm{e}^{-y_\ell x} \mathbb{P}(A_{n,j}(x/a_{n,j}))dx\right)\right| \\ &\leq \sum_{\ell=1}^{\varsigma} \sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} \left|\mathbb{E}\left(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i)}\right) - \left(1 - \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \int_0^\infty y_\ell \mathrm{e}^{-y_\ell x} \mathbb{P}(A_{n,j}(x/a_{n,j}))dx\right)\right| \\ &\leq \sum_{\ell=1}^{\varsigma} \sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1} |Err| \leq \sum_{i=1}^{k_n} |Err| \\ &\leq 2 \sum_{i=1}^{k_n} \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \sum_{r=j+1}^{\mathcal{L}_{H,n,i}-1} \mathbb{P}\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_{n,r}\}\right) + \sum_{i=1}^{k_n} \int_0^\infty y_\ell \mathrm{e}^{-y_\ell x} \delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}}(x/a_n)dx \\ &\to 0\end{aligned}$$

as $n \to \infty$ by $Д'_q(u_{n,i})^*$ and $ULC_q(u_{n,i})$, so it follows that

$$\begin{aligned}
\prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\mathbb{E}\Big(\mathrm{e}^{-y_\ell a_n \mathscr{A}_n(J_i)}\Big) &\sim \prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\left(1-\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\int_0^\infty y_\ell \mathrm{e}^{-y_\ell x}\mathbb{P}(A_{n,j}(x/a_{n,j}))dx\right)\\
&\sim \prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\left(1-\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\mathbb{P}\left(Q^{(0)}_{q,n,j}\right)\int_0^\infty y_\ell \mathrm{e}^{-y_\ell x}(1-\pi(x))dx\right)\\
&\sim \prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\left(1-\theta\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})\left(1-\pi(0)-\int_0^\infty \mathrm{e}^{-y_\ell x}d\pi(x)\right)\right)\\
&= \prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\left(1-\theta(1-\phi(y_\ell))\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})\right)\\
&\sim \mathrm{e}^{-\sum_{\ell=1}^{\varsigma}\sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\theta(1-\phi(y_\ell))\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})}
\end{aligned}$$

where ϕ is the Laplace transform of π, and since we have, by (2.3),

$$\left|\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})-\frac{\ell_{H,n,i}}{n}\tau\right| \le \left|\sum_{j=0}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})-\frac{\mathcal{L}_{H,n,i}}{n}\tau\right|+\left|\sum_{j=0}^{\mathcal{L}_{H,n,i-1}-1}\bar{F}(u_{n,j})-\frac{\mathcal{L}_{H,n,i-1}}{n}\tau\right| \to 0$$

then, by (3.9),

$$\frac{\tau}{n}\sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\ell_{H,n,i} \sim \frac{\tau}{n}.n\left|\frac{1}{\tau}I_\ell\right| = |I_\ell|$$

and

$$\left|\sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})-|I_\ell|\right| \le \sum_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\left|\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1}\bar{F}(u_{n,j})-\frac{\tau}{n}\ell_{H,n,i}\right| \to 0.$$

We conclude that

$$\mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma}y_\ell a_n\mathscr{A}_n(v_nI_\ell)}\right) \sim \prod_{\ell=1}^{\varsigma}\prod_{i=i_\ell}^{i_\ell+\mathscr{S}_\ell-1}\mathbb{E}\Big(\mathrm{e}^{-y_\ell a_n\mathscr{A}_n(J_i)}\Big) \sim \mathrm{e}^{-\theta\sum_{\ell=1}^{\varsigma}(1-\phi(y_\ell))|I_\ell|} = \mathbb{E}\left(\mathrm{e}^{-\sum_{\ell=1}^{\varsigma}y_\ell A(I_\ell)}\right)$$

where A is a compound Poisson process of intensity θ and multiplicity d.f. π.

□

4. Convergence of the REPP

When the mark function m_n defined in (2.7) counts the number of exceedances then our atomic random measure A_n is actually a REPP as the one considered in [FFT13], namely, $A_n(J) = \sum_{j\in v_nJ\cap\mathbb{N}_0}\mathbf{1}_{X_j>u_{n,j}}$. Suppose that we have a system with decay of correlations against L^1 and ζ is the only global maximum of φ, which is a periodic point of prime period p, then for all large

$n \in \mathbb{N}$, $\{X_j > u_{n,j}\} \cap \{X_{j+k} > u_{n,j+k}\} \neq \emptyset$ if and only if k is a multiple of p. So, we can set

$$U^{(\kappa)}_{p,n,j} = \{X_j > u_{n,j}, X_{j+p} > u_{n,j+p}, \dots, X_{j+\kappa p} > u_{n,j+\kappa p}\}$$
$$Q^{(\kappa)}_{p,n,j} = \{X_j > u_{n,j}, X_{j+p} > u_{n,j+p}, \dots, X_{j+\kappa p} \leq u_{n,j+\kappa p}\}$$

and we note the following:

$$m_n(I_\kappa) > x \Leftrightarrow \kappa \geq \lfloor x \rfloor \qquad R^{(\kappa)}_{n,j}(x) = \begin{cases} Q^{(\kappa)}_{p,n,j} & \text{if } \kappa \geq \lfloor x \rfloor \\ \emptyset & \text{if } \kappa < \lfloor x \rfloor \end{cases}$$
$$B_{n,j}(x) = \bigcup_{\kappa=\lfloor x \rfloor}^{\infty} Q^{(\kappa)}_{p,n,j} \cup U^{(\infty)}_{p,n,j} = U^{(\lfloor x \rfloor)}_{p,n,j}$$
$$A_{n,j}(x) = U^{(\lfloor x \rfloor)}_{p,n,j} \cap \bigcap_{\ell=1}^{p} (U^{(\lfloor x \rfloor)}_{p,n,j+\ell})^c = Q^{(\lfloor x \rfloor)}_{p,n,j}$$

Let π be a multiplicity distribution satisfying (2.13), with normalising factors $a_{n,j} = 1$. Then, for every $H \in \mathbb{N}$, $i = 0, 1, \dots, Hn - 1$ and $x \geq 0$,

$$\pi(x) \sim 1 - \frac{\mathbb{P}(A_{n,j}(x))}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)} = 1 - \frac{\mathbb{P}\left(Q^{(\lfloor x \rfloor)}_{p,n,j}\right)}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)}$$

and

$$E(\pi) = \int x d\pi \sim \int x d\left(1 - \frac{\mathbb{P}\left(Q^{(\lfloor x \rfloor)}_{p,n,j}\right)}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)}\right) = \sum_{\kappa=1}^{\infty} \kappa \frac{\mathbb{P}\left(Q^{(\kappa-1)}_{p,n,j}\right) - \mathbb{P}\left(Q^{(\kappa)}_{p,n,j}\right)}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)} =$$
$$= \sum_{j=1}^{\infty} \sum_{\kappa=j}^{\infty} \frac{\mathbb{P}\left(Q^{(\kappa-1)}_{p,n,j}\right) - \mathbb{P}\left(Q^{(\kappa)}_{p,n,j}\right)}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)} = \sum_{j=1}^{\infty} \frac{\mathbb{P}\left(Q^{(j-1)}_{p,n,j}\right)}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)} = \frac{\mathbb{P}\left(U^{(0)}_{p,n,j}\right)}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)} = \frac{\bar{F}(u_{n,j})}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)}$$

so that $\mathbb{P}\left(Q^{(0)}_{p,n,j}\right) \sim \frac{1}{E(\pi)} \bar{F}(u_{n,j})$ and we conclude that the Extremal Index θ is given by $\frac{1}{E(\pi)}$ which is equivalent to say that the first moment of the multiplicity distribution π is $1/\theta$.

Additionally, suppose that ζ is a repelling point, which means that we have backward contraction implying that $U^{(\infty)}_{j,n,i} = \{\zeta\}$ and implying that there exists $0 < \theta < 1$ so that $U^{(\kappa)}_{p,n,j}$ is a ball around ζ with

$$\mathbb{P}\left(U^{(\kappa)}_{p,n,j}\right) \sim (1-\theta)^\kappa \mathbb{P}(X_j > u_{n,j})$$

Then, we recover the main result in [FFT13], which states that A_n converges in distribution to a compound Poisson process of intensity θ and geometric multiplicity distribution π. In fact, in this case we have $\mathbb{P}\left(Q^{(\kappa)}_{p,n,j}\right) = \mathbb{P}\left(U^{(\kappa)}_{p,n,j}\right) - \mathbb{P}\left(U^{(\kappa+1)}_{p,n,j}\right) \sim \theta(1-\theta)^\kappa \mathbb{P}(X_j > u_{n,j})$ and the result follows from observing that $\mathbb{P}\left(Q^{(0)}_{p,n,j}\right) \sim \theta \mathbb{P}(X_j > u_{n,j}) = \theta \bar{F}(u_{n,j})$, which means that θ is the

Extremal Index, and that, for every $H \in \mathbb{N}$, $i = 0, 1, \ldots, Hn - 1$,

$$\frac{\mathbb{P}(A_{n,j}(x))}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)} = \frac{\mathbb{P}\left(Q^{(\lfloor x \rfloor)}_{p,n,j}\right)}{\mathbb{P}\left(Q^{(0)}_{p,n,j}\right)} \sim (1-\theta)^{\lfloor x \rfloor} = 1 - \pi(x)$$

where $\pi(x) = 1 - (1-\theta)^{\lfloor x \rfloor}$ is the cumulative distribution function of a geometric distribution of parameter θ, that is, $\pi(x) = \sum_{\kappa \leq x, \kappa \in \mathbb{N}} \theta(1-\theta)^{\kappa - 1}$.

5. Application to sequential systems: an example of uniformly expanding map

In this section we will give a detailed analysis of the application of the general result obtained in Section 2 to a particular sequential system. It is constructed with β transformations, although it can be generalised to other examples of sequential systems presented in [FFV17, Section 3] after making the necessary adaptations.

Consider the family of maps on the unit circle $S^1 = [0, 1]$, with the identification $0 \sim 1$, given by $T_\beta(x) = \beta x \bmod 1$ for $\beta > 1$. Note that for many such β, we have that $T_\beta(1) \neq 1$ and, by the identification $0 \sim 1$, this means that T_β as a map on S^1 is not continuous at $\zeta = 0 \sim 1$. For simplicity we assume that $T_\beta(0) = 0$ but consider that the orbit of 1 is still defined to be $T_\beta(1), T^2_\beta(1), \ldots$ although, strictly speaking, $1 \sim 0$ should be considered a fixed point. In what follows m denotes Lebesgue measure on $[0, 1]$.

Theorem 5.A. *Consider an unperturbed map T_β corresponding to some $\beta = \beta_0 > 1 + c$, with invariant absolutely continuous probability $\mu = \mu_\beta$. Consider a sequential system acting on the unit circle and given by $\mathcal{T}_n = T_n \circ \cdots \circ T_1$, where $T_i = T_{\beta_i}$, for all $i = 1, \ldots, n$ and $|\beta_n - \beta| \leq n^{-\xi}$ holds for some $\xi > 1$. Let $X_1, X_2, \ldots$ be defined by (2.1), where the observable function φ achieves a global maximum at a chosen periodic point ζ of prime period p [1] (we allow $\varphi(\zeta) = +\infty$), being of following form:*

$$\varphi(x) = g\big(dist(x, \zeta)\big), \tag{5.1}$$

where the function $g : [0, +\infty) \to \mathbb{R} \cup \{+\infty\}$ achieves its global maximum at 0 ($g(0)$ may be $+\infty$); is a strictly decreasing homeomorphism $g : V \to W$ in a neighbourhood V of 0; and has one of the following three types of behaviour:

Type 1: there exists some strictly positive function $h : W \to \mathbb{R}$ such that for all $y \in \mathbb{R}$

$$\lim_{s \to g(0)} \frac{g^{-1}(s + yh(s))}{g^{-1}(s)} = e^{-y}; \tag{5.2}$$

Type 2: $g(0) = +\infty$ and there exists $\beta > 0$ such that for all $y > 0$

$$\lim_{s \to +\infty} \frac{g^{-1}(sy)}{g^{-1}(s)} = y^{-\beta}; \tag{5.3}$$

Type 3: $g(0) = D < +\infty$ and there exists $\gamma > 0$ such that for all $y > 0$

$$\lim_{s \to 0} \frac{g^{-1}(D - sy)}{g^{-1}(D - s)} = y^{\gamma}. \tag{5.4}$$

Remark 5.1. The different types of g imply that the distribution of X_0 falls in the domain of attraction for maxima of the Gumbel, Fréchet and Weibull distributions, respectively.

[1] $T^p_\beta(\zeta) = \zeta$ and p is the minimum integer with such property

Let $(u_n)_{n\in\mathbb{N}}$ be such that $n\mu(X_0 > u_n) \to \tau$, as $n \to \infty$ for some $\tau \geq 0$.

Then, the POT MREPP a_nA_n converges in distribution to a compound Poisson process with intensity θ given by

$$\theta = \begin{cases} 1-\beta^{-p}, \text{when the orbit of } \zeta \text{ by } T_\beta \text{ never hits } 0\sim 1 \\ \frac{d\mu}{dm}(0)(1-\beta^{-1}) + \frac{d\mu}{dm}(1)(1-\beta^{-p}), \text{when } \zeta = 0 \sim 1 \end{cases} \tag{5.5}$$

and multiplicity distribution

$$\pi(x) = \begin{cases} 1-e^{-x}, \text{when } g \text{ is of type 1 and } a_n = h(u_n)^{-1} \\ 1-(1+x)^{-\beta}, \text{when } g \text{ is of type 2 and } a_n = u_n^{-1} \\ 1-(1-x)^{\gamma}, \text{when } g \text{ is of type 3 and } a_n = (D-u_n)^{-1} \end{cases} \tag{5.6}$$

and, for $a_n = h(u_n)^{-1}$ the AOT MREPP a_nA_n converges in distribution to a compound Poisson process with the same intensity θ as above and multiplicity d.f. π given by

$$\pi(x) = 1 - \lim_{n\to\infty} h_{\kappa(u_n, q(u_n)x)}(x) \tag{5.7}$$

where $g_{\kappa,u}(x) = \sum_{i=0}^{\kappa}(g(M^i x) - u)$ with $M = \beta^p$; $\kappa = \kappa(u,x)$ is the only integer such that $x \in \left[g_{\kappa,u}\left(\frac{g^{-1}(u)}{M^\kappa}\right), g_{\kappa,u}\left(\frac{g^{-1}(u)}{M^{\kappa+1}}\right)\right)$; and h_k is a strictly monotone homeomorphism h_k such that

$$\lim_{u\to g(0)} \frac{g_{\kappa,u}(g^{-1}(u)h_\kappa(x))}{h(u)} = x. \tag{5.8}$$

Remark 5.2. Examples of each one of the three types are as follows: $g(x) = -\log x$ (in this case (5.2) is easily verified with $h \equiv 1$), $g(x) = x^{-1/\alpha}$ for some $\alpha > 0$ (condition (5.3) is verified with $\beta = \alpha$) and $g(x) = D - x^{1/\alpha}$ for some $D \in \mathbb{R}$ and $\alpha > 0$ (condition (5.4) is verified with $\gamma = \alpha$). For these examples, the multiplicity d.f. of the compound Poisson process associated to the AOT MREPP a_nA_n can be computed as shown in the following table:

Examples of $g(x)$	Respective distribution $\pi(x)$
$-\log(x)$	$1-(\sqrt{M})^{-\lfloor\frac{\sqrt{1+8x/\log M}-1}{2}\rfloor}\mathrm{e}^{-\frac{x}{\lfloor\frac{\sqrt{1+8x/\log M}-1}{2}\rfloor+1}}$
$x^{-1/\alpha}$	$1-\left(\frac{1-M^{-1/\alpha}}{1-M^{-(\kappa(x)+1)/\alpha}}\right)^{-\alpha}(\kappa(x)+1+x)^{-\alpha}$ where $\kappa = \kappa(x)$ is the only integer such that $\frac{M^{\kappa/\alpha}-M^{-1/\alpha}}{1-M^{-1/\alpha}} \leq \kappa+1+x < \frac{M^{(\kappa+1)/\alpha}-1}{1-M^{-1/\alpha}}$
$D-x^{1/\alpha}$	$1-\left(\frac{1-M^{1/\alpha}}{1-M^{(\kappa(x)+1)/\alpha}}\right)^{\alpha}(\kappa(x)+1-x)^{\alpha}$ where $\kappa = \kappa(x)$ is the only integer such that $\frac{1-M^{-(\kappa+1)/\alpha}}{M^{1/\alpha}-1} < \kappa+1-x \leq \frac{M^{1/\alpha}-M^{-\kappa/\alpha}}{M^{1/\alpha}-1}$

Remark 5.3. We point out that in this example we take $u_{n,i} = u_n$, where $(u_n)_{n\in\mathbb{N}}$ satisfies $n\mu(X_0 > u_n) \to \tau$, as $n \to \infty$ for some $\tau > 0$, where μ is the invariant measure of the original map T_β.

5.1. **Preliminaries.** As we said above, we let μ denote the invariant measure of the original map T_β and let $h = \frac{d\mu}{dm}$ be its density. In what follows, let $U_n = \{X_0 > u_n\}$.

We will assume throughout this subsection the existance of some $\xi > 1$ such that

$$|\beta_n - \beta| \leq \frac{1}{n^\xi}. \tag{5.9}$$

Also let $0 < \gamma < 1$ be such that $\gamma\xi > 1$. In what follows P denotes the Perron-Fröbenius transfer operator associated to the unperturbed map T_β with respect to the reference Lebesgue measure m, that is, the operator defined by the duality relation

$$\int Pf\ g\ dm\ =\ \int f\ g \circ T_\beta\ dm, \quad \text{for all } f \in L^1_m,\ g \in L^\infty_m.$$

Recall that $\Pi_i = P_i \circ \ldots \circ P_1$, where P_i is the transfer operator associated to $T_i = T_{\beta_i}$, while P^i is the corresponding concatenation for the unperturbed map T_β. Note that by [CR07, Lemma 3.10], we have

$$\left\| \Pi_i(g) - \int g dm\ h \right\|_1 \leq C_1 \frac{\log i}{i^\xi} \|g\|_{BV}. \tag{5.10}$$

For any measurable set $A \subset [0,1]$, we have

$$\begin{aligned} m(\mathcal{T}_i^{-1}(A)) &= \int \mathbf{1}_A \circ T_i \circ \ldots \circ T_1 dm = \int \mathbf{1}_A \Pi_i(1) dm \\ &= \int \mathbf{1}_A h dm + \int \mathbf{1}_A (\Pi_i(1) - h) dm. \end{aligned}$$

By (5.10), if $i \geq \lfloor n^\gamma \rfloor$ (recall that $\gamma\xi > 1$) then we have $\int |\Pi_i(1) - h| dm \leq C_1 \frac{\log i}{i^\xi} = o(n^{-1})$, which allows us to write:

$$m(\mathcal{T}_i^{-1}(A)) = \mu(A) + o(n^{-1}). \tag{5.11}$$

5.1.1. *Verification of condition* (2.3). We want to show that $\sum_{i=0}^{h_n-1} m(X_i > u_n) = \frac{h_n}{n}\tau + o(1)$ for any unbounded increasing sequence of positive integers $h_n \leq Hn$.

We begin with the following lemma.

Lemma 5.4. *We have that*

$$\sum_{i=0}^{h_n-1} \int_{U_n} P^i(1)\, dm = \frac{h_n}{n}\tau + o(1).$$

Proof. By hypothesis, for all $i \in \mathbb{N}$ and $g \in BV$ we have $P^i(g) = h \int g \cdot h\, dm + Q^i(g)$, where $\|Q^i(g)\|_\infty \leq \alpha^i \|g\|_{BV}$, for some $\alpha < 1$. Then we can write:

$$\begin{aligned} \sum_{i=0}^{h_n-1} \int_{U_n} P^i(1) dm &= \sum_{i=0}^{h_n-1} \int h \left(\int 1 \cdot h dm \right) \mathbf{1}_{U_n} dm + \sum_{i=0}^{h_n-1} \int Q^i(1) \mathbf{1}_{U_n} dm \\ &= \sum_{i=0}^{h_n-1} \int_{U_n} h dm + \sum_{i=0}^{h_n-1} \int Q^i(1) \mathbf{1}_{U_n} dm \\ &= \frac{h_n}{n} n\mu(U_n) + \sum_{i=0}^{h_n-1} \int Q^i(1) \mathbf{1}_{U_n} dm. \end{aligned}$$

The result will follow once we show that the second term on the right goes to 0, as $n \to \infty$. This follows easily because

$$\sum_{i=0}^{h_n-1} \int Q^i(1)\mathbf{1}_{U_n} dm \leq \sum_{i=0}^{h_n-1} \alpha^i \int \mathbf{1}_{U_n} dm = \frac{1-\alpha^{h_n}}{1-\alpha} m(U_n) \xrightarrow[n\to\infty]{} 0.$$

□

Since

$$\sum_{i=0}^{h_n-1} m(X_i > u_n) = \sum_{i=0}^{h_n-1} \int_{U_n} \Pi_i(1) dm = \sum_{i=0}^{h_n-1} \int_{U_n} P^i(1) dm + \sum_{i=0}^{h_n-1} \int_{U_n} \Pi_i(1) - P^i(1) dm,$$

then condition (2.5) holds once we prove that the second term on the right goes to 0 as $n \to \infty$.

Let $\varepsilon > 0$ be arbitrary. Since $\xi > 1$ then $\sum_{i\geq 0} \frac{\log i}{i^\xi} < \infty$, so there exists $N \geq \lfloor n^\gamma \rfloor$ such that $C_1 \sum_{i\geq N} \frac{\log i}{i^\xi} < \varepsilon/2$.

On the other hand, using the Lasota-Yorke inequalities (see [FFV17, Section 3]) for both Π and P, we have that there exists some $C > 0$ such that $|\Pi_i(1) - P^i(1)| \leq C$, for all $i \in \mathbb{N}$. Let n be sufficiently large so that $CNm(U_n) < \varepsilon/2$. Then

$$\begin{aligned}\left|\sum_{i=0}^{h_n-1} \int_{U_n} \Pi_i(1) - P^i(1) dm\right| &\leq \sum_{i=0}^{N-1} \int_{U_n} |\Pi_i(1) - P^i(1)| dm + \sum_{i=N}^{\infty} \int_{U_n} |\Pi_i(1) - P^i(1)| dm \\ &\leq CNm(U_n) + C_1 \sum_{i\geq N} \frac{\log i}{i^\xi} < \varepsilon/2 + \varepsilon/2 = \varepsilon.\end{aligned}$$

5.2. **Verification of condition** $Д_q(u_{n,i})^*$**.** We will use the following proposition, proved in [FFV17, Section 3].

Proposition 5.5. *Let $\phi \in BV$ and $\psi \in L^1(m)$. Then for the β transformations $T_n = T_{\beta_n}$ we have that*

$$\left|\int \phi \circ \mathcal{T}_i \psi \circ \mathcal{T}_{i+t} dm - \int \phi \circ \mathcal{T}_i dm \int \psi \circ \mathcal{T}_{i+t} dm\right| \leq B\lambda^t \|\phi\|_{BV} \|\psi\|_1,$$

for some $\lambda < 1$ and $B > 0$ independent of ϕ and ψ.

Remark 5.6. As it can be seen in [CR07, Section 3], Proposition 5.5 holds for any sequence $T_{\beta_1}, T_{\beta_2}, \ldots$ of β transformations and not necessarily only for the ones that satisfy condition (5.9).

Condition $Д_q(u_{n,i})^*$ follows from Proposition 5.5 by taking for each $i \leq Hn - 1$,

$$\phi_i = \mathbf{1}_{D_{n,i}(x_1)} \text{ and } \psi_i = \mathbf{1}_{\bigcap_{j=2}^{s} \{\mathscr{A}_n(I_j - i - t) \leq x_j\}},$$

where for every $j \leq Hn - 1$ we define

$$D_{n,j}(x) := B_{n,0}(x) \cap \bigcap_{\ell=1}^{q} (T_{j+\ell} \circ \ldots \circ T_{j+1})^{-1}(B_{n,0}(x))^c. \tag{5.12}$$

Since we assume that (5.9) holds, there exists a constant $C > 0$ depending on x_1 but not on i such that $\|\phi_i\|_{BV} < C$. Moreover, it is clear that $\|\psi_i\|_1 \leq 1$. Hence,

$$\left| m\left(A_{n,i}(x_1) \cap \bigcap_{j=2}^{\varsigma} \{\mathscr{A}_n(I_j) \leq x_j\} \right) - m\left(A_{n,i}(x_1)\right) m\left(\bigcap_{j=2}^{\varsigma} \{\mathscr{A}_n(I_j) \leq x_j\} \right) \right|$$
$$= \left| \int \phi_i \circ \mathcal{T}_i \psi_i \circ \mathcal{T}_{i+t} dm - \int \phi_i \circ \mathcal{T}_i dm \int \psi_i \circ \mathcal{T}_{i+t} dm \right| \leq \text{const } \lambda^t.$$

Thus, if we take $\gamma_i(n,t) = \text{const}\lambda^t$ and $t_n^* = (\log n)^2$, condition $Д_q(u_{n,i})^*$ is trivially satisfied.

5.3. **Verification of condition $Д_q'(u_{n,i})^*$.** We start by noting that we may neglect the first $\lfloor n^\gamma \rfloor$ random variables of the process $X_0, X_1, \ldots$, where γ is such that $\gamma\xi > 1$, for ξ given as in (5.9).

In fact, by the Uniform Doeblin-Fortet-Lasota-Yorke inequality (DFLY) used in [FFV17, Section 3], we have

$$m(\mathscr{A}_n([\lfloor n^\gamma \rfloor, n)) \leq x) - m(\mathscr{A}_n([0,n)) \leq x) = m(\mathscr{A}_n([0, \lfloor n^\gamma \rfloor)) > 0) \leq \sum_{i=0}^{\lfloor n^\gamma \rfloor - 1} m(X_i > u_n)$$
$$= \sum_{i=0}^{\lfloor n^\gamma \rfloor - 1} \int \mathbf{1}_{U_n} \Pi_i(1) dm \leq C_0 n^\gamma m(U_n) \xrightarrow[n\to\infty]{} 0.$$

This way, we simply disregard the $\lfloor n^\gamma \rfloor$ random variables of $X_0, X_1, \ldots$ and start the blocking procedure, described in Section 3.1, in $X_{\lfloor n^\gamma \rfloor}$ by taking $\mathcal{L}_{H,n,0} = \lfloor n^\gamma \rfloor$. We split the remaining $n - \lfloor n^\gamma \rfloor$ random variables into k_n blocks as described in Section 3.1. Our goal is to show that

$$S_n' := \sum_{i=1}^{k_n} \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} \sum_{r>j}^{\mathcal{L}_{H,n,i}-1} m\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_n\}\right) + \sum_{j=\mathcal{L}_{H,n,k_n}}^{Hn-1} \sum_{r>j}^{Hn-1} m\left(Q_{q,n,j}^{(0)} \cap \{X_r > u_n\}\right)$$

goes to 0.

We define for some $j, n, q \in \mathbb{N}_0$,

$$R_{q,n,j} := \min\left\{ r \in \mathbb{N} : \; Q_{q,n,j}^{(0)} \cap \{X_{j+r} > u_n\} \neq \emptyset \right\},$$
$$\tilde{R}_{q,n} := \min\{R_{q,n,j}, \; j = \lfloor n^\gamma \rfloor, \ldots, Hn-1\},$$
$$L_n := \max\{\ell_{H,n,i}, \; i = 1, \ldots, k_n\}$$
$$\tilde{L}_n := \max\{L_n, Hn - \mathcal{L}_{H,n,k_n}\}.$$

We have

$$S_n' \leq \sum_{j=\lfloor n^\gamma \rfloor}^{Hn-1} \sum_{r \geq R_{q,n,j}}^{\tilde{L}_n} m\left(Q_{q,n,j}^{(0)} \cap \{X_{j+r} > u_n\}\}\right) = \sum_{j=\lfloor n^\gamma \rfloor}^{Hn-1} \sum_{r \geq R_{q,n,i}}^{\tilde{L}_n} \int \mathbf{1}_{D_{q,n,j}} \circ \mathcal{T}_j \cdot \mathbf{1}_{U_n} \circ \mathcal{T}_{j+r} \, dm,$$

where for every $j \leq Hn - 1$ we define

$$D_{q,n,j} := U_n \cap \bigcap_{\ell=1}^{q} (T_{j+\ell} \circ \ldots \circ T_{j+1})^{-1}(U_n)^c. \tag{5.13}$$

Using Proposition 5.5, with $\phi = \mathbf{1}_{D_{q,n,j}}$ and $\psi = \mathbf{1}_{U_n}$, and the adjoint property of the operators, it follows that

$$\int \mathbf{1}_{D_{q,n,j}} \circ \mathcal{T}_j \cdot \mathbf{1}_{U_n} \circ \mathcal{T}_{j+r}\, dm \leq \int \mathbf{1}_{D_{q,n,j}} \Pi_j(1) dm \int \mathbf{1}_{U_n} \Pi_{j+r}(1) dm + B\lambda^r \|\mathbf{1}_{D_{q,n,j}}\|_{BV} \|\mathbf{1}_{U_n}\|_1.$$

Using (DFLY), we have

$$\int \mathbf{1}_{D_{q,n,j}} \circ \mathcal{T}_j \cdot \mathbf{1}_{U_n} \circ \mathcal{T}_{j+r}\, dm \leq C_0^2 m(U_n)^2 + BC_2\lambda^r m(U_n)$$

for some $C_2 > 0$ (independent of n) such that $\|\mathbf{1}_{D_{q,n,j}}\|_{BV} \leq C_2$. Hence,

$$\begin{aligned} S_n' &\leq \sum_{j=\lfloor n^\gamma \rfloor}^{Hn-1} \sum_{r \geq R_{q,n,i}}^{\tilde{L}_n} \left(C_0^2 m(U_n)^2 + BC_2\lambda^r m(U_n) \right) \leq C_0^2 Hn\tilde{L}_n m(U_n)^2 + BC_2 m(U_n) Hn \sum_{r \geq \tilde{R}_{q,n}}^{\tilde{L}_n} \lambda^r \\ &\leq C_0^2 Hn\tilde{L}_n m(U_n)^2 + BC_2 m(U_n) Hn \lambda^{\tilde{R}_{q,n}} \frac{1}{1-\lambda}. \end{aligned}$$

Now we show that $\tilde{L}_n = o(n)$. To see this, observe that each $\ell_{H,n,i}$ is defined, in this case, by the largest integer ℓ such that

$$\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i-1}+\ell-1} m(X_j > u_n) \leq \frac{1}{k_n} \sum_{j=\lfloor n^\gamma \rfloor}^{Hn-1} m(X_j > u_n).$$

Using (5.11), it follows that

$$\ell_{H,n,i}\mu(U_n)(1+o(1)) \leq \frac{Hn - \lfloor n^\gamma \rfloor}{k_n} \mu(U_n)(1+o(1)).$$

On the other hand, by definition of $\ell_{H,n,i}$ we must have

$$\sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i-1}+\ell_{H,n,i}-1} m(X_j > U_n) > \frac{1}{k_n} \sum_{j=\lfloor n^\gamma \rfloor}^{Hn-1} m(X_j > u_n) - m(X_{\mathcal{L}_{H,n,i-1}+\ell_{H,n,i}} > u_n).$$

Using (5.11) again, we have

$$\ell_{H,n,i}\mu(U_n)(1+o(1)) > \frac{Hn - \lfloor n^\gamma \rfloor}{k_n} \mu(U_n)(1+o(1)) - \mu(U_n)(1+o(1)).$$

Together with the previous inequality, we have

$$\ell_{H,n,i} = \frac{Hn - \lfloor n^\gamma \rfloor}{k_n}(1+o(1)) = o(n) \tag{5.14}$$

for every $i = 1, \ldots, k_n$ and

$$Hn - \mathcal{L}_{H,n,k_n} = Hn - \lfloor n^\gamma \rfloor - \sum_{i=1,\ldots,k_n} \ell_{H,n,i} = (Hn - \lfloor n^\gamma \rfloor) o(1) = o(n)$$

so $\tilde{L}_n = o(n)$ follows at once. Using this estimate, the fact that $\lim_{n\to\infty} n\mu(U_n) = \tau$ and $h \in BV$, we have $C_0^2 Hn\tilde{L}_n m(U_n)^2 \to 0$.

In order to prove that $Д_q'(u_{n,i})^*$ holds, we need to show that $\tilde{R}_{q,n} \to \infty$, as $n \to \infty$. To do that we consider two cases, whether the orbit of ζ hits 1 or not.

We will consider that the maps T_i, for all $i \in \mathbb{N}_0$, are defined in S^1 by using the usual identification $0 \sim 1$. Observe that the only possible point of discontinuity of such maps is $0 \sim 1$. Moreover, $\lim_{x\to 0^+} T_i(x) = 0$ and $\lim_{x\to 1^-} T_i(x) = \beta_i - \lfloor \beta_i \rfloor$.

5.3.1. *The orbit of ζ by the unperturbed T_β map does not hit 1.* We mean that for all $j \in \mathbb{N}_0$ we have $T^j(\zeta) \neq 1$.

We take $q = p$, where $p \in \mathbb{N}$ is such that $T^p(\zeta) = \zeta$ and $T^j(\zeta) \neq \zeta$ for all $j < p$. Let

$$\varepsilon_n := |\beta_{\lfloor n^\gamma \rfloor} - \beta|. \tag{5.15}$$

By (5.9) and choice of γ, we have that $\varepsilon_n = o(n^{-1})$. Also let $\delta > 0$, be such that $B_\delta(\zeta)$ is contained on a domain of injectivity of all T_i, with $i \geq \lfloor n^\gamma \rfloor$.

Let $J \in \mathbb{N}$ be chosen. Using a continuity argument, we can show that there exists $C := C(J, q) > 0$ such that

$$\operatorname{dist}(T_{i+j} \circ \ldots \circ T_{i+1}(\zeta), T^j(\zeta)) < C\varepsilon_n, \text{ for all } j = 1, \ldots, J$$

and moreover $U_n \cap T_{i+j} \circ \ldots \circ T_{i+1}(U_n) = \emptyset$, for all $j \leq J$ such that $j/q - \lfloor j/q \rfloor > 0$.

We want to check that if $x \in Q^{(0)}_{q,n,i}$ for some $i \geq \lfloor n^\gamma \rfloor$, *i.e.*, $\mathcal{T}_i(x) \in D_{q,n,i}$, then $X_{i+j}(x) \leq u_n$, for all $j = 1, \ldots, J$. By the assumptions above, we only need to check the latter for all $j = 1, \ldots, J$ such that $j/q - \lfloor j/q \rfloor = 0$, *i.e.*, for all $j = sq$, where $s = 1, \ldots, \lfloor J/q \rfloor$.

By definition of $Q^{(0)}_{q,n,i}$ the statement is clearly true when $s = 1$. Now, we consider $s > 1$ and let $x \in Q^{(0)}_{q,n,i}$. We have

$$\operatorname{dist}(\mathcal{T}_{i+sq}(x), T_{i+sq} \circ \ldots \circ T_{i+q+1}(\zeta)) > (\beta - \varepsilon_n)^{(s-1)q} \operatorname{dist}(\mathcal{T}_{i+q}(x), \zeta).$$

On the other hand,

$$\operatorname{dist}(T_{i+sq} \circ \ldots \circ T_{i+q+1}(\zeta), \zeta) \leq C\varepsilon_n.$$

Hence,

$$\begin{aligned}
\operatorname{dist}(\mathcal{T}_{i+sq}(x), \zeta) &\geq \operatorname{dist}(\mathcal{T}_{i+sq}(x), T_{i+sq} \circ \ldots \circ T_{i+q+1}(\zeta)) - \operatorname{dist}(T_{i+sq} \circ \ldots \circ T_{i+q+1}(\zeta), \zeta) \\
&\geq (\beta - \varepsilon_n)^{(s-1)q} \operatorname{dist}(\mathcal{T}_{i+q}(x), \zeta) - C\varepsilon_n \\
&\geq (\beta - \varepsilon_n)^{(s-1)q} \frac{m(U_n)}{2} - C\varepsilon_n, \text{ since } x \in Q^{(0)}_{q,n,i} \Rightarrow X_{i+q}(x) \leq u_n \Leftrightarrow \mathcal{T}_{i+q}(x) \notin U_n \\
&> \frac{m(U_n)}{2}, \text{ for } n \text{ sufficiently large, since } \varepsilon_n = o(n^{-1}).
\end{aligned}$$

This shows that $\mathcal{T}_{i+sq}(x) \notin U_n$, which means that $X_{i+sq}(x) \leq u_n$.

5.3.2. $\zeta = 0 \sim 1$. In this case we proceed in the same way as in [AFV15, Section 3.3], which basically corresponds to considering two versions of the same point: $\zeta^+ = 0$ and $\zeta^- = 1$. Note that ζ^+ is a fixed point for all maps considered and ζ^- is periodic of prime period p.

As the previous case, we take $q = p$. We observe that $D_{q,n,i}$ has two connected components, one to the right of 0 and the other to the left of 1, where none of the two points belongs to the set. Let $J \in \mathbb{N}$ be fixed as before. A continuity argument as the one used before allows us to show that the points of the components of $D_{q,n,i}$ do not return to U_n before J iterates, also. Note that, the maps are orientation preserving so there is no switching as described in [AFV15, Section 3.3].

5.4. **Verification of condition** (2.12)**.** Similarly to the previous condition, we disregard the first $\lfloor n^\gamma \rfloor$ random variables of $X_0, X_1, \dots$ and start the blocking procedure in $X_{\lfloor n^\gamma \rfloor}$ by taking $\mathcal{L}_{H,n,0} = \lfloor n^\gamma \rfloor$. We want to show that

$$\lim_{n\to\infty} \max_{i=1,\dots,k_n} \left\{ \left| \theta k_n \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} m(X_j > u_n) - k_n \sum_{j=\mathcal{L}_{H,n,i-1}}^{\mathcal{L}_{H,n,i}-1} m\left(Q^{(0)}_{q,n,j}\right) \right| \right\} = 0.$$

Let ε_n be defined as in (5.15) and let δ_n be such that $U_n = B_{\delta_n}(\zeta)$. For simplicity, we assume that we are using the usual Riemannian metric so that we have a symmetry of the balls, which means that $|U_n| = m(U_n) = 2\delta_n$.

We also assume that ζ is a periodic point of prime period p with respect to the unperturbed map $T = T_\beta$ and the orbit of ζ does not hit $0 \sim 1$. In this case, we take $\theta = 1 - \beta^{-q}$ with $q = p$ and check (2.12).

Using a continuity argument we can show that there exists $C := C(J, q) > 0$ such that

$$\operatorname{dist}(T_{i+q} \circ \dots \circ T_{i+1}(\zeta), \zeta) < C\varepsilon_n.$$

We define two points ξ_u and ξ_l of $B_{\delta_n}(\zeta)$ on the same side with respect to ζ such that $\operatorname{dist}(\xi_u, \zeta) = (\beta - \varepsilon_n)^{-q}\delta_n + C\varepsilon_n$ and $\operatorname{dist}(\xi_l, \zeta) = (\beta + \varepsilon_n)^{-q}\delta_n - (\beta + \varepsilon_n)^{-q}C\varepsilon_n$. Recall that for all $i \geq \lfloor n^\gamma \rfloor$, we have that $(\beta - \varepsilon_n)^q \leq \beta_{i+1} \cdot \dots \cdot \beta_{i+q} \leq (\beta + \varepsilon_n)^q$.

Since we are composing β transformations, then for all $i \geq \lfloor n^\gamma \rfloor$, we have

$$\operatorname{dist}(T_{i+q} \circ \dots \circ T_i(\xi_u), T_{i+q} \circ \dots \circ T_i(\zeta)) \geq \delta_n + (\beta - \varepsilon_n)^q C\varepsilon_n.$$

Using the triangle inequality it follows that

$$\operatorname{dist}(T_{i+q} \circ \dots \circ T_{i+1}(\xi_u), \zeta) \geq \delta_n.$$

Similarly, $\operatorname{dist}(T_{i+q} \circ \dots \circ T_{i+1}(\xi_l), T_{i+q} \circ \dots \circ T_{i+1}(\zeta)) \leq \delta_n - C\varepsilon_n$ and

$$\operatorname{dist}(T_{i+q} \circ \dots \circ T_{i+1}(\xi_l), \zeta) \leq \delta_n.$$

If we assume that both ξ_u and ξ_l are on the right hand side with respect to ζ and ξ_u^* and ξ_l^* are the corresponding points on the left hand side of ζ, then

$$(\zeta - \delta_n, \xi_u^*] \cup [\xi_u, \zeta + \delta_n) \subset D_{q,n,i} \subset (\zeta - \delta_n, \xi_l^*] \cup [\xi_l, \zeta + \delta_n).$$

Hence,

$$\delta_n - (\beta - \varepsilon_n)^{-q}\delta_n - C\varepsilon_n \leq \frac{1}{2} m(D_{q,n,i}) \leq \delta_n - (\beta + \varepsilon_n)^{-q}\delta_n + (\beta + \varepsilon_n)^{-q}C\varepsilon_n.$$

Since $\varepsilon_n = o(n^{-1}) = o(\delta_n)$ then we easily get

$$\lim_{n\to\infty} \frac{m(D_{q,n,i})}{m(U_n)} = 1 - \beta^{-q}.$$

Observe that by (5.11), $m(Q^{(0)}_{q,n,i}) = m(\mathcal{T}_i^{-1}(D_{q,n,i})) = \mu(D_{q,n,i}) + o(n^{-1})$ and $m(X_i > u_n) = \mu(U_n) + o(n^{-1})$. Hence, we have that

$$\lim_{n\to\infty} \frac{m(Q^{(0)}_{q,n,i})}{m(X_i > u_n)} = \lim_{n\to\infty} \frac{\mu(D_{q,n,i})}{\mu(U_n)}.$$

The density $\frac{d\mu}{dm}$, which can be found in [Par60, Theorem 2], is sufficiently regular so that, as in [FFT15, Section 7.3], one can see that

$$\lim_{n\to\infty}\frac{\mu(D_{q,n,i})}{\mu(U_n)}=\lim_{n\to\infty}\frac{m(D_{q,n,i})}{m(U_n)}.$$

It follows that

$$\lim_{n\to\infty}\frac{m(Q^{(0)}_{q,n,i})}{m(X_i>u_n)}=1-\beta^{-q}.$$

Since, as we have seen in (5.14), we can write $\ell_{H,n,i}=\frac{Hn}{k_n}(1+o(1))$, then the previous equation can easily be used to prove that condition (2.12) holds, with $\theta=1-\beta^{-q}$.

For the case $\zeta=0\sim 1$ the argument will follow similarly, although we have to take into account the fact that the density is discontinuous at $0\sim 1$. By [Par60] we have that

$$\frac{d\mu}{dm}(x)=\frac{1}{M(\beta)}\sum_{x<T^n(1)}\frac{1}{\beta^n},$$

where $M(\beta):=\int_0^1\sum_{x<T^n(1)}\frac{1}{\beta^n}dm$. In this case, we have $\theta=\frac{d\mu}{dm}(0)(1-\beta^{-1})+\frac{d\mu}{dm}(1)(1-\beta^{-q})$.

5.5. **Verification of condition** (2.13)**.** Once again, we disregard the first $\lfloor n^\gamma\rfloor$ random variables of $X_0,X_1,\ldots$. We want to show that

$$\lim_{n\to\infty}\max_{j=\lfloor n^\gamma\rfloor,\ldots,Hn-1}\left\{\left|\frac{m(A_{n,j}(x/a_n))}{m\left(Q^{(0)}_{q,n,j}\right)}-(1-\pi(x))\right|\right\}=0.$$

Observing that by (5.11), $m(B_{n,i}(x))=m(\mathcal{T}_i^{-1}(B_{n,0}(x)))=\mu(B_{n,0}(x))+o(n^{-1})$ and using an argument similar to the one of the previous condition, we have

$$\lim_{n\to\infty}\frac{m(A_{n,i}(x))}{m(B_{n,i}(x))}=\lim_{n\to\infty}\frac{\mu(D_{n,i}(x))}{\mu(B_{n,0}(x))}=\lim_{n\to\infty}\frac{m(D_{n,i}(x))}{m(B_{n,0}(x))}=\theta$$

where

$$D_{n,j}:=\mathcal{T}_j^{-1}(Q^{(0)}_{q,n,j})=U_n\cap\bigcap_{\ell=1}^{q}(T_{j+\ell}\circ\ldots\circ T_{j+1})^{-1}(U_n)^c \tag{5.16}$$

and with the same θ as before. Hence,

$$\lim_{n\to\infty}\frac{m(A_{n,i}(x))}{m(Q^{(0)}_{q,n,i})}=\lim_{n\to\infty}\frac{\theta m(B_{n,i}(x))}{\theta m(X_i>u_n)}=\lim_{n\to\infty}\frac{m(B_{n,i}(x))}{m(X_i>u_n)}$$

Let $\tilde{B}_{n,i}(x)$ be the set $B_{n,i}(x)$ associated to the unperturbed dynamical system given by $\mathcal{T}_n=(T_\beta)^n$ and $\tilde{X}_i$ the corresponding unperturbed random variables. Using a continuity argument we can show that $m(B_{n,i}(x))\sim m(\tilde{B}_{n,i}(x))$ and $m(X_i>u_n)\sim m(\tilde{X}_i>u_n)$, so that

$$\lim_{n\to\infty}\frac{m(B_{n,i}(x))}{m(X_i>u_n)}=\lim_{n\to\infty}\frac{m(\tilde{B}_{n,i}(x))}{m(\tilde{X}_i>u_n)}=\lim_{n\to\infty}\frac{\mu(\tilde{B}_{n,i}(x))}{\mu(\tilde{X}_i>u_n)}=\lim_{n\to\infty}\frac{\mu(\tilde{B}_{n,0}(x))}{\mu(U_n)}$$

For this unperturbed stationary process, it has been proved in [FFMa18, Section 3] that $\lim_{n\to\infty}\frac{\mu(\tilde{B}_{n,0}(x/a_n))}{\mu(U_n)}=1-\pi(x)$, where $\pi(x)$ is the distribution given in (5.6) for the POT MREPP a_nA_n and given in

(5.7) for the AOT MREPP a_nA_n. So, $\lim_{n\to\infty}\frac{m(A_{n,i}(x/a_n))}{m(Q^{(0)}_{q,n,i})}=1-\pi(x)$ for that same distribution $\pi(x)$ and for any $i=\lfloor n^\gamma\rfloor,\dots,Hn-1$. Hence, (2.13) follows at once.

5.6. **Verification of condition $ULC_q(u_{n,i})$.** We want to see that, for all $H\in\mathbb{N}$ and $y>0$,

$$\lim_{n\to\infty}\sum_{i=1}^{k_n}\int_0^\infty \mathrm{e}^{-yx}\delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}}(x/a_n)dx=0,$$

$$\lim_{n\to\infty}\int_0^\infty \mathrm{e}^{-x}\delta_{n,\mathcal{L}_{H,n,k_n},Hn-\mathcal{L}_{H,n,k_n}}(x/a_n)dx=0,$$

$$\text{and}\quad \lim_{n\to\infty}\sum_{i=1}^{k_n}\int_0^\infty \mathrm{e}^{-yx}\delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}-t_{H,n,i}}(x/a_n)dx=0$$

where a_n is as in (2.13) and $\delta_{n,s,\ell}(x)$ as in (2.14). Then, for all $x\in\mathbb{R}_0^+$,

$$\begin{aligned}\delta_{n,s,\ell}(x)&\le\sum_{\kappa=1}^{\lfloor\ell/q\rfloor}\sum_{j=s+\ell-\kappa q}^{s+\ell-1}m\left(Q^{(\kappa)}_{q,n,j}\right)+\sum_{j=s}^{s+\ell-1}\sum_{\kappa>\lfloor\ell/q\rfloor}m\left(Q^{(\kappa)}_{q,n,j}\right)+\sum_{j=1}^{q}m\left(U^{(0)}_{q,n,s+\ell-j}\right)\\&\le\sum_{\kappa=1}^{\infty}\sum_{j=s+\ell-\kappa q}^{s+\ell-1}m\left(Q^{(\kappa)}_{q,n,j}\right)+\sum_{j=1}^{q}m\left(U^{(0)}_{q,n,s+\ell-j}\right)\end{aligned}$$

hence for all $x\in\mathbb{R}_0^+$ and $y\in\mathbb{R}^+$, we have

$$\sum_{i=1}^{k_n}\int_0^\infty \mathrm{e}^{-yx}\delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}}(x/a_n)dx\le\frac{1}{y}\sum_{i=1}^{k_n}\left(\sum_{\kappa=1}^{\infty}\sum_{j=\mathcal{L}_{H,n,i}-\kappa q}^{\mathcal{L}_{H,n,i}-1}m\left(Q^{(\kappa)}_{q,n,j}\right)+\sum_{j=1}^{q}m\left(U^{(0)}_{q,n,\mathcal{L}_{H,n,i}-j}\right)\right),$$

$$\int_0^\infty \mathrm{e}^{-x}\delta_{n,\mathcal{L}_{H,n,k_n},Hn-\mathcal{L}_{H,n,k_n}}(x/a_n)dx\le\sum_{\kappa=1}^{\infty}\sum_{j=Hn-\kappa q}^{Hn-1}m\left(Q^{(\kappa)}_{q,n,j}\right)+\sum_{j=1}^{q}m\left(U^{(0)}_{q,n,Hn-j}\right)$$

$$\text{and}\quad \sum_{i=1}^{k_n}\int_0^\infty \mathrm{e}^{-yx}\delta_{n,\mathcal{L}_{H,n,i-1},\ell_{H,n,i}-t_{H,n,i}}(x/a_n)dx$$

$$\le\frac{1}{y}\sum_{i=1}^{k_n}\left(\sum_{\kappa=1}^{\infty}\sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}-\kappa q}^{\mathcal{L}_{H,n,i}-t_{H,n,i}-1}m\left(Q^{(\kappa)}_{q,n,j}\right)+\sum_{j=1}^{q}m\left(U^{(0)}_{q,n,\mathcal{L}_{H,n,i}-t_{H,n,i}-j}\right)\right).$$

Let $\tilde{Q}^{(\kappa)}_{q,n,j}$ and $\tilde{U}^{(0)}_{q,n,j}$ be the corresponding sets $Q^{(\kappa)}_{q,n,j}$ and $U^{(0)}_{q,n,j}$ associated to the unperturbed dynamical system given by $\mathcal{T}_n=(T_\beta)^n$. Using a continuity argument we can show that $m\left(Q^{(\kappa)}_{q,n,j}\right)\sim m\left(\tilde{Q}^{(\kappa)}_{q,n,j}\right)$ and $m\left(U^{(0)}_{q,n,j}\right)\sim m\left(\tilde{U}^{(0)}_{q,n,j}\right)$. For this unperturbed stationary process, it has been proved in [FFMa18, Section 3] that

$$m\left(\tilde{Q}^{(\kappa)}_{q,n,j}\right)\sim\theta(1-\theta)^\kappa m\left(\tilde{U}^{(0)}_{q,n,j}\right).$$

so we have $m\left(Q^{(\kappa)}_{q,n,j}\right) \sim \theta(1-\theta)^\kappa m\left(U^{(0)}_{q,n,j}\right)$. Additionally, using (5.11) (once again neglecting the first $\lfloor n^\gamma \rfloor$ random variables), $m\left(U^{(0)}_{q,n,j}\right) = m(\mathcal{T}_j^{-1}(U_n)) \sim \mu(U_n) \sim m(U_n)$, so $m\left(Q^{(\kappa)}_{q,n,j}\right) \sim \theta(1-\theta)^\kappa m(U_n)$ and, by (2.9),

$$\sum_{i=1}^{k_n}\left(\sum_{\kappa=1}^{\infty}\sum_{j=\mathcal{L}_{H,n,i}-\kappa q}^{\mathcal{L}_{H,n,i}-1} m\left(Q^{(\kappa)}_{q,n,j}\right) + \sum_{j=1}^{q} m\left(U^{(0)}_{q,n,\mathcal{L}_{H,n,i}-j}\right)\right)$$

$$\sim k_n\left(\sum_{\kappa=1}^{\infty}\kappa q\theta(1-\theta)^\kappa m(U_n) + qm(U_n)\right) = \frac{k_n q}{\theta} m(U_n) \xrightarrow[n\to\infty]{} 0.$$

Similarly,

$$\sum_{\kappa=1}^{\infty}\sum_{j=Hn-\kappa q}^{Hn-1} m\left(Q^{(\kappa)}_{q,n,j}\right) + \sum_{j=1}^{q} m\left(U^{(0)}_{q,n,Hn-j}\right) \sim \frac{q}{\theta} m(U_n) \xrightarrow[n\to\infty]{} 0$$

and

$$\sum_{i=1}^{k_n}\left(\sum_{\kappa=1}^{\infty}\sum_{j=\mathcal{L}_{H,n,i}-t_{H,n,i}-\kappa q}^{\mathcal{L}_{H,n,i}-t_{H,n,i}-1} m\left(Q^{(\kappa)}_{q,n,j}\right) + \sum_{j=1}^{q} m\left(U^{(0)}_{q,n,\mathcal{L}_{H,n,i}-t_{H,n,i}-j}\right)\right)$$

$$\sim k_n\left(\sum_{\kappa=1}^{\infty}\kappa q\theta(1-\theta)^\kappa m(U_n) + qm(U_n)\right) = \frac{k_n q}{\theta} m(U_n) \xrightarrow[n\to\infty]{} 0.$$

6. Random dynamical systems

We now give another example of a non-stationary system in the form of a *fibred dynamical system* constructed by taking Lasota-Yorke maps on the fibers; we refer in particular to the paper [DFGTV18].

Let us consider the unit interval $I = [0,1]$, endowed with the Borel σ-algebra $\mathcal{B}$ and the Lebesgue measure m. Furthermore, let

$$\mathrm{var}(g) = \inf_{h=g(\mathrm{mod}\ m)} \sup_{0=s_0<s_1<\ldots<s_n=1} \sum_{k=1}^{n} |h(s_k) - h(s_{k-1})|.$$

the variation of the function $g \in L^1(m)$. We define $BV(I,m)$ (sometimes shortened in BV), as the Banach space with respect to the norm

$$\|h\|_{BV} = \mathrm{var}(h) + \|h\|_1.$$

For a piecewise C^2 function $f : [0,1] \to [0,1]$, set $\delta(f) = \operatorname{ess\,inf}_{x\in[0,1]}|f'|$ and let $N(f)$ denote the number of intervals of monotonicity of f. Then let $(\Omega, \mathcal{F}, \mathbb{Q})$ be a probability space and let $\sigma : \Omega \to \Omega$ be an invertible $\mathbb{Q}$-preserving transformation. We will assume that $\mathbb{Q}$ is ergodic. Consider now a measurable map $\omega \mapsto f_\omega$, $\omega \in \Omega$ of piecewise C^2 maps on $[0,1]$ defined as above such that

$$N := \sup_{\omega\in\Omega} N(f_\omega) < \infty,\ \delta := \inf_{\omega\in\Omega} \delta(f_\omega) > 1, \quad \text{and } D := \sup_{\omega\in\Omega} |f''_\omega|_\infty < \infty. \tag{6.1}$$

and such that the map $(\omega, x) \mapsto (\mathcal{P}_\omega H(\omega, \cdot))(x)$ is $\mathbb{Q} \times m$-measurable for any $\mathbb{Q} \times m$ measurable function H with $H(\omega, \cdot) \in L^1(m)$ for a.e. $\omega \in \Omega$, where $\mathcal{P}_\omega$ denotes the transfer (Perron-Fröbenius) operator associated to f_ω.

For next purposes, we need two more assumption.

- First we ask that the following uniform covering condition holds: for every subinterval $J \subset I, \exists k = k(J) \in \mathbb{N}$ such that for a.e. $\omega \in \Omega$, $f_{\sigma^{k-1}\omega} \circ \cdots \circ f_\omega(J) = I$.
- Then we require the existence of $N \in \mathbb{N}$ such that for each $a > 0$ and any sufficiently large $n \in \mathbb{N}$, there is $c > 0$ such that

$$\operatorname{ess\,inf} \mathcal{P}_\omega^{Nn} h \geqslant c/2 \|h\|_1, \quad \text{for every } h \in C_a \text{ and a.e. } \omega \in \Omega,$$

  where $C_a := \{\phi \in BV : \phi \geqslant 0 \text{ and } \operatorname{var}(\phi) \leqslant a \int \phi \, dm\}$.
  This cone-type condition will guarantee that the density h_ω constructed below is strictly positive, namely

$$\operatorname{ess\,inf} h_\omega \geqslant c/2, \text{ for a.e. } \omega \in \Omega. \tag{6.2}$$

The next step is to introduce the probability governing the extreme value distributions. First of all we can associate to our collection of mappings on I, $f_\omega \colon I \to I$, $\omega \in \Omega$ the skew product transformation $\tau \colon \Omega \times I \to \Omega \times I$ defined by

$$\tau(\omega, x) = (\sigma\omega, f_\omega(x)). \tag{6.3}$$

The preceding bunch of assumptions on the maps f_ω, allows us to show that there exist a unique measurable and nonnegative function $h_\omega \colon \Omega \times I \to \mathbb{R}$ with the property that $h_\omega := h(\omega, \cdot) \in BV$, $\int h_\omega \, dm = 1$, $\mathcal{L}_\omega(h_\omega) = h_{\sigma\omega}$ for a.e. $\omega \in \Omega$ and

$$\operatorname{ess\,sup}_{\omega \in \Omega} \|h_\omega\|_{BV} < \infty. \tag{6.4}$$

If we now define a probability measure μ on $\Omega \times I$ by

$$\mu(A \times B) = \int_{A \times B} h_\omega \, d(\mathbb{Q} \times m), \quad \text{for } A \in \mathcal{F} \text{ and } B \in \mathcal{B}, \tag{6.5}$$

then it follows that μ is invariant with respect to τ. Furthermore, μ is obviously absolutely continuous with respect to $\mathbb{Q} \times m$ and is the only measure with these properties.

Let us now consider for any $\omega \in \Omega$ the measures μ_ω on the measurable space $(I, \mathcal{B})$, defined by $d\mu_\omega = h_\omega dm$. We recall here two important properties of these measures. First, the so-called *equivariant property*: $f_\omega^* \mu_\omega = \mu_{\sigma\omega}$. Second, the *disintegration* of μ on the marginal $\mathbb{Q}$: if A is any measurable set in $\mathcal{F} \times \mathcal{B}$, and $A_\omega = \{x : (\omega, x) \in A\}$, the *section* at ω, then $\mu(A) = \int \mu_\omega(A_\omega) d\mathbb{Q}(\omega)$.

The conditional (or sample) measure μ_ω will constitute the probability underlying our random processes, which we called $\mathbb{P}$ in the preceding sections.

After this preparatory work we can now state the decay of correlations result which will be used later on. Let μ_ω be, as above, the measure on X given by $d\mu_\omega = h_\omega dm$ for $\omega \in \Omega$. Then there exists $K > 0$ and $\rho \in (0, 1)$ such that

$$\left| \int \phi \, \psi \circ f_\omega^n \, d\mu_\omega - \int \phi \, d\mu_\omega \cdot \int \psi \, d\mu_{\sigma^n \omega} \right| \leqslant K \rho^n \|\psi\|_1 \cdot \|\phi\|_{BV}, \tag{6.6}$$

for $n \geqslant 0$, $\psi \in L^1(m)$ and $\phi \in BV(X,m)$; $\|\cdot\|_1$ denotes the L^1 norm with respect to m.[2]

We now choose $\Omega = Y^{\mathbb{Z}}$, where $Y = (1, \cdots, l)$ is a finite alphabet with l letters. We associate to each letter a map satisfying the requirements given above: we call them *random Lasota-Yorke* maps. The map σ will therefore be the bilateral shift and $\mathbb{Q}$ any ergodic shift-invariant non-atomic ergodic probability measure, for instance, and it is the choice we do here, a Bernoulli measure with weights $p_1, \cdots, p_l$.

We now consider the process given by $X_k := \phi \circ f^k_\omega, k \in \mathbb{N}$, where $f^k_\omega := f_{\omega_k} \circ \cdots \circ f_{\omega_1}$, being $\omega_j \in Y, j = 1, \cdots k$, the first k symbols of the word ω. The function $\phi : I \to \mathbb{R} \cup \{\pm\infty\}$ achieves a global maximum at $z \in I$ (we allow $\phi(z) = +\infty$), being of the following form: $\phi(x) = g(\mathrm{dist}(x,z))$, where $g : [0, +\infty) \to \mathbb{R} \cup \{+\infty\}$ achieves its global maximum at 0 ($g(0)$ may be $+\infty$,) and g is a strictly decreasing bijection in a neighborhood of 0. Finally g assumes one of three types of behavior which we recalled in the statement of Theorem 5.A. We now introduce the marginal measure μ_I on I as: $\mu_I(B) = \int_\Omega \mu_\omega(B) d\mathbb{Q}(\omega)$, with B a measurable subset of I. As in [FFV17] we consider all the boundary levels equal $u_{n,i} = u_n, i = 1, \cdots, n-1$, where u_n is determined by the marginal measure μ_I so that

$$\mu_n = \inf\{u \in \mathbb{R} : \mu_I(\{x \in I : \phi(x) \leqslant u\}) \geqslant 1 - \frac{\tau}{n}\}, \tag{6.7}$$

for some $\tau > 0$. With this choice and by Lemma 9 of [RSV14] we have

$$\sum_{i=0}^{Hn-1} \mu_{\sigma^i \omega}(\{x \in I : \phi(x) > u_n\}) \to \tau, \text{ as } n \to \infty, \tag{6.8}$$

which is our equation (2.3) for the fibred systems. From now on we will set $U_n := \{x \in I : \phi(x) > u_n\}$ which, by the choice of the function g, is an open neighborhood of the point z.

Condition $Д_q(u_{n,i})^*$ with $\mathbb{P} = \mu_\omega$ can now be worked out easily thanks to the decay of correlations (6.6), which takes care of observables given by characteristic functions, see the function $\psi \in L^1(m)$ in (6.6). We defer for the details to the second part of Proposition 4.3 in [FFV17] which is the same as in the present context.

We now go the condition $Д'_q(u_{n,i})^*$. We should first of all elaborate about the choice of the target point z. Since we have finitely many maps f_k each of which with finitely many branches, we could choose the point z on a set of full m measure in such a way that it will not intersect the preimages of any order of any of the maps $f_1, \cdots, f_m$. We should also remember that the statement on the convergence in distribution for the extreme value law should hold for $\mathbb{Q}$-almost all choice of ω defining the sample measure μ_ω. This will be useful in the following *periodicity* considerations, which will allow us to choose $q = 0$ in the conditions $Д_q(u_{n,i})^*$ and $Д'_q(u_{n,i})^*$ above. We begin to notice that three situations can occur:

- For a given ω, the point z will never come back to itself, namely $f^k_\omega z \neq z, \forall k \geqslant 1$.
- For a given ω there are finitely many blocks of periodicity, namely we have finitely many sequences of type $\omega_{i_1} \cdots \omega_{i_L} \in \omega$ for which $f_{\omega_{i_L}} \cdots f_{\omega_{i_1}} z = z$.
- For a given ω there are countably many blocks of periodicity like those described in the preceding item.

[2]The result in [DFGTV18], Lemma 4, is stated in a different manner. It requires ψ in $L^\infty(m)$. Since the density h_ω is in $L^\infty(m)$ too as an element of $BV(X,m)$, and moreover is essentially bounded uniformly in ω by (6.4), we get the $\|\cdot\|_1$ norm on the right hand side of (6.6).

We begin to observe that the set of the words with infinitely many blocks of periodicity has measure zero. We therefore treat now the words with finitely many blocks of periodicity, the situation in the first item being included in that one. Having fixed such an ω, call n_ω the last time $f_\omega^{n_\omega} z = z$. The proof follows now closely that in section 4.3.1 on the [FFV17] paper to which we defer for the details. We now point out the main differences arising in our framework.

- First of all we use the quenched decay of correlations established in (6.6) applied to the same observable $\mathbf{1}_{U_n}$. This will produce two asymptotic terms $\mu_{\sigma^i\omega}(U_n)$ and $\mu_{\sigma^j\omega}(U_n), j > i$, and the exponential error term containing the Lebesgue measure $m(U_n)$.
- The measure of $\mu_{\sigma^i\omega}(U_n)$ will appear in a sum ranging from 1 to n and therefore it will converge to τ by (6.8).
- The other measure should be expressed in terms of the Lebesgue measure m in order to compare it with the error term and to establish bounds from below and from above for the quantity $\mathcal{L}_{H,n} := \max\{\mathcal{L}_{H,n,i}, i = 1, \cdots, k_n\}$. Thanks to (6.2) and (6.4), we have that there exists two constants c_1 and c_2 such that for $\mathbb{Q}$-almost any $\omega \in \Omega$ we have that
$$c_1 m(U_n) \leqslant \mu_{\sigma^i\omega}(U_n) \leqslant c_2 m(U_n), \ \forall i \geqslant 1.$$
- We now come to the main difference with the analogous proof in section 4.3.1 in [FFV17]. We have to prove that if $f_\omega^i(x) \in U_n$, then $f_\omega^j(x) \in U_n$, for the next time with j growing to infinity. We already put n_ω the last time $f_\omega^{n_\omega} z = z$. If we now fix $J \in \mathbb{N}$, then $f_\omega^{n_\omega+k} z, \ k = 1, \cdots, J$, will never return to z. Since we are composing finitely many maps, there will be an $\varepsilon > 0$, such that $\forall \omega \in \Omega$ and $k = 1, \cdots, J$ we have $\mathrm{dist}(f_\omega^{n_{\omega+k}} z, z) > \varepsilon$. Call $\overline{n}$ the integer such that $\mathrm{diameter}(U_{\overline{n}}) < \frac{\varepsilon}{4}\delta^{-J}$ and $U_{\overline{n}}$ does not intersect the preimages up to order J of the family of maps $f_k, k = 1, \cdots, J$. If we now take $n > \max\{n_\omega, \overline{n}\}$, we have that $\forall x \in U_n$ $\mathrm{dist}(f_\omega^J x, z) > \frac{\varepsilon}{2}$.

We are left with the verification of conditions 2.12 and 2.13. By (6.8) and the definition of $Q_{0,n,j}^{(0)} = \{\phi \circ f_\omega^j > u_n\}$, we see immediately that $\theta = 1$. The computation of 2.13 follows closely that in the proof of Theorem 3.A in [FFMa18]; we give the details for the type-1 observable $g = -\log x$, for which $h = 1$. We are reduced to estimate the ratio $\frac{\mu_{\sigma^j}(X_0 > u_n + x)}{\mu_{\sigma^j}(X_0 > u_n)}$, where $X_0(\cdot) = -\log \mathrm{dist}(\cdot, z)$ and z is chosen m-almost everywhere. We have

$$\frac{\mu_{\sigma^j}(X_0 > u_n + x)}{\mu_{\sigma^j}(X_0 > u_n)} = \frac{m(B(z, e^{-u_n - x}))}{m(B(z, e^{-u_n}))} \frac{m(B(z, e^{-u_n})) \int_{B(z, e^{-u_n - x})} h_{\sigma^j\omega} dm}{m(B(z, e^{-u_n - x})) \int_{B(z, e^{-u_n})} h_{\sigma^j\omega} dm},$$

where $B(z, v)$ denotes a ball of center z and radius v. In the limit of large n the ratio on the right hand side of the preceding equality goes to 1 by Lebesgue's differentiation theorem, while the first ratio on the left hand side goes to e^{-x}. This gives the desired result with the probability distribution $\pi = 1 - e^{-x}$. By generalizing we easily get the equivalent of Theorem 5.A in our case

Proposition 6.1. *For the random fibred system constructed above and having chosen the observable $\phi(x) = g(dist(x, z))$, where g has one of the three forms given in the statement of Theorem 5.A and z is chosen m-almost everywhere, the POT and AOT MREPP $a_n A_n$ both converge in distribution to a compound Poisson distribution process with intensity $\theta = 1$ and multiplicity distribution*

$$\pi(x) = \begin{cases} 1 - e^{-x}, \text{when } g \text{ is of type 1 and } a_n = h(u_n)^{-1} \\ 1 - (1 + x)^{-\beta}, \text{when } g \text{ is of type 2 and } a_n = u_n^{-1} \\ 1 - (1 - x)^{\gamma}, \text{when } g \text{ is of type 3 and } a_n = (D - u_n)^{-1} \end{cases} \tag{6.9}$$

References

[AFV15] Hale Aytaç, Jorge Milhazes Freitas, and Sandro Vaienti, *Laws of rare events for deterministic and random dynamical systems*, Trans. Amer. Math. Soc. **367** (2015), no. 11, 8229–8278. MR 3391915

[AHN+15] Romain Aimino, Huyi Hu, Matthew Nicol, Andrei Török, and Sandro Vaienti, *Polynomial loss of memory for maps of the interval with a neutral fixed point*, Discrete Contin. Dyn. Syst. **35** (2015), no. 3, 793–806. MR 3277171

[BB84] Daniel Berend and Vitaly Bergelson, *Ergodic and mixing sequences of transformations*, Ergodic Theory Dynam. Systems **4** (1984), no. 3, 353–366. MR 776873 (86i:28017)

[CFM+18] Théophile Caby, Davide Faranda, Giorgio Mantica, Sandro Vaienti, and Pascal Yiou, *Generalized dimensions, large deviations and the distribution of rare events*, Preprint arXiv:1812.00036, November 2018.

[CHM91] Michael R. Chernick, Tailen Hsing, and William P. McCormick, *Calculating the extremal index for a class of stationary sequences*, Adv. in Appl. Probab. **23** (1991), no. 4, 835–850. MR MR1133731 (93c:60073)

[Col01] P. Collet, *Statistics of closest return for some non-uniformly hyperbolic systems*, Ergodic Theory Dynam. Systems **21** (2001), no. 2, 401–420. MR MR1827111 (2002a:37038)

[CR07] Jean-Pierre Conze and Albert Raugi, *Limit theorems for sequential expanding dynamical systems on* $[0, 1]$, Ergodic theory and related fields, Contemp. Math., vol. 430, Amer. Math. Soc., Providence, RI, 2007, pp. 89–121. MR 2331327 (2008j:37077)

[DFGTV18] D. Dragi*v*cević, G. Froyland, C. González-Tokman, and S. Vaienti, *Almost sure invariance principle for random piecewise expanding maps*, Nonlinearity **31** (2018), no. 5, 2252–2280. MR 3816673

[FACM+19] Davide Faranda, M. Carmen Alvarez-Castro, Gabriele Messori, David Rodrigues, and Pascal Yiou, *The hammam effect or how a warm ocean enhances large scale atmospheric predictability*, Nature Communications **10** (2019), no. 1, 1316.

[FFMa18] Ana Cristina Moreira Freitas, Jorge Milhazes Freitas, and Mário Magalhães, *Convergence of marked point processes of excesses for dynamical systems*, J. Eur. Math. Soc. (JEMS) **20** (2018), no. 9, 2131–2179. MR 3836843

[FFRS18] Ana Cristina Moreira Freitas, Jorge Milhazes Freitas, Fagner Rodrigues, and Jorge Soares, *Rare events for cantor target sets*, Preprint arXiv:1903.07200, March 2018.

[FFT10] Ana Cristina Moreira Freitas, Jorge Milhazes Freitas, and Mike Todd, *Hitting time statistics and extreme value theory*, Probab. Theory Related Fields **147** (2010), no. 3-4, 675–710. MR 2639719 (2011g:37015)

[FFT11] ———, *Extreme value laws in dynamical systems for non-smooth observations*, J. Stat. Phys. **142** (2011), no. 1, 108–126. MR 2749711 (2012a:60149)

[FFT12] ———, *The extremal index, hitting time statistics and periodicity*, Adv. Math. **231** (2012), no. 5, 2626–2665. MR 2970462

[FFT13] ———, *The compound Poisson limit ruling periodic extreme behaviour of non-uniformly hyperbolic dynamics*, Comm. Math. Phys. **321** (2013), no. 2, 483–527. MR 3063917

[FFT15] ———, *Speed of convergence for laws of rare events and escape rates*, Stochastic Process. Appl. **125** (2015), no. 4, 1653–1687. MR 3310360

[FFV17] Ana Cristina Moreira Freitas, Jorge Milhazes Freitas, and Sandro Vaienti, *Extreme Value Laws for non stationary processes generated by sequential and random dynamical systems*, Ann. Inst. Henri Poincaré Probab. Stat. **53** (2017), no. 3, 1341–1370. MR 3689970

[FFV18] ———, *Extreme value laws for sequences of intermittent maps*, Proc. Amer. Math. Soc. **146** (2018), no. 5, 2103–2116. MR 3767361

[HNTV17] Nicolai Haydn, Matthew Nicol, Andrew Török, and Sandro Vaienti, *Almost sure invariance principle for sequential and non-stationary dynamical systems*, Trans. Amer. Math. Soc. **369** (2017), no. 8, 5293–5316. MR 3646763

[Hüs83] Jürg Hüsler, *Asymptotic approximation of crossing probabilities of random sequences*, Z. Wahrsch. Verw. Gebiete **63** (1983), no. 2, 257–270. MR 701529 (84j:60044)

[Hüs86] ———, *Extreme values of nonstationary random sequences*, J. Appl. Probab. **23** (1986), no. 4, 937–950. MR 867190 (88e:60030)

[Kal86] Olav Kallenberg, *Random measures*, fourth ed., Akademie-Verlag, Berlin, 1986. MR 854102 (87k:60137)

[Kar91] Alan F. Karr, *Point processes and their statistical inference*, second ed., Probability: Pure and Applied, vol. 7, Marcel Dekker, Inc., New York, 1991. MR 1113698

[Kif88] Yuri Kifer, *Random perturbations of dynamical systems*, Progress in Probability and Statistics, vol. 16, Birkhäuser Boston, Inc., Boston, MA, 1988. MR 1015933 (91e:58159)

[Lea91] M. R. Leadbetter, *On a basis for "peaks over threshold" modeling*, Statist. Probab. Lett. **12** (1991), no. 4, 357–362. MR 1131065 (93a:60077)

[LFF+16] Valerio Lucarini, Davide Faranda, Ana Cristina Moreira Freitas, Jorge Milhazes Freitas, Mark Holland, Tobias Kuna, Matthew Nicol, and Sandro Vaienti, *Extremes and recurrence in dynamical systems*, Pure and Applied Mathematics: A Wiley Series of Texts, Monographs and Tracts, Wiley, Hoboken, NJ, 2016.

[LFTV12] Valerio Lucarini, Davide Faranda, Giorgio Turchetti, and Sandro Vaienti, *Extreme value theory for singular measures*, Chaos **22** (2012), no. 2, –.

[MCB+18] Gabriele Messori, Rodrigo Caballero, Freddy Bouchet, Davide Faranda, Richard Grotjahn, Nili Harnik, Steve Jewson, Joaquim G. Pinto, Gwendal Rivière, Tim Woollings, and Pascal Yiou, *An interdisciplinary approach to the study of extreme weather events: Large-scale atmospheric controls and insights from dynamical systems theory and statistical mechanics*, Bulletin of the American Meteorological Society **99** (2018), no. 5, ES81–ES85.

[MCF17] Gabriele Messori, Rodrigo Caballero, and Davide Faranda, *A dynamical systems approach to studying midlatitude weather extremes*, Geophysical Research Letters **44** (2017), no. 7, 3346–3354.

[NTV18] Matthew Nicol, Andrew Török, and Sandro Vaienti, *Central limit theorems for sequential and random intermittent dynamical systems*, Ergodic Theory Dynam. Systems **38** (2018), no. 3, 1127–1153. MR 3784257

[Par60] W. Parry, *On the β-expansions of real numbers*, Acta Math. Acad. Sci. Hungar. **11** (1960), 401–416. MR 0142719 (26 #288)

[Rou14] Jérôme Rousseau, *Hitting time statistics for observations of dynamical systems*, Nonlinearity **27** (2014), no. 9, 2377–2392. MR 3266858

[RSV14] Jérôme Rousseau, Benoit Saussol, and Paulo Varandas, *Exponential law for random subshifts of finite type*, Stochastic Process. Appl. **124** (2014), no. 10, 3260–3276. MR 3231619

[RT15] Jérôme Rousseau and Mike Todd, *Hitting times and periodicity in random dynamics*, J. Stat. Phys. **161** (2015), no. 1, 131–150. MR 3392511

[SKF+16] E. W. Saw, D. Kuzzay, D. Faranda, A. Guittonneau, F. Daviaud, C. Wiertel-Gasquet, V. Padilla, and B. Dubrulle, *Experimental characterization of extreme events of inertial dissipation in a turbulent swirling flow*, Nature Communications **7** (2016), 12466 EP –.

Ana Cristina Moreira Freitas, Centro de Matemática & Faculdade de Economia da Universidade do Porto, Rua Dr. Roberto Frias, 4200-464 Porto, Portugal

Email address: amoreira@fep.up.pt

URL: http://www.fep.up.pt/docentes/amoreira/

Jorge Milhazes Freitas, Centro de Matemática & Faculdade de Ciências da Universidade do Porto, Rua do Campo Alegre 687, 4169-007 Porto, Portugal

Email address: jmfreita@fc.up.pt

URL: http://www.fc.up.pt/pessoas/jmfreita/

Mário Magalhães, Centro de Matemática da Universidade do Porto, Rua do Campo Alegre 687, 4169-007 Porto, Portugal

Email address: mdmagalhaes@fc.up.pt

Sandro Vaienti, Aix Marseille Université, Université de Toulon, CNRS, CPT, 13009 Marseille, France.

Email address: vaienti@cpt.univ-mrs.fr

URL: http://www.cpt.univ-mrs.fr/~vaienti/