

# No evidence for behavioural syndrome and genetic basis for three personality traits in a wild bird population

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1	No evidence for behavioural syndrome and genetic basis for three
2	personality traits in a wild bird population
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23	Keywords: aggressiveness, between- and within-individual correlations, boldness, collared
24	flycatcher, Ficedula albicollis, heritability, neophobia, parental effects, pedigree, quantitative
25	genetics, repeatability

#### 26 ABSTRACT

27

Personality traits and their correlations have been shown to be linked with life history 28 29 strategies and fitness in various species. Between-individual correlations (i.e. behavioural syndromes) between personality traits can affect the evolutionary responses of these traits to 30 environmental variation. Understanding the genetic and ecological determinants of 31 personality traits and their interactions as behavioural syndromes in the wild is thus needed to 32 shed light on the mechanisms shaping their evolution. Partitioning the observed (co)variance 33 in these traits, however, requires large numbers of repeated behavioural measures on many 34 35 individuals of known relatedness level. In the absence of such data, it is thus often assumed that phenotypic (co)variances inform about (1) underlying between-individual (co)variances 36 (i.e. ignoring within-individual (co)variances) and (2) underlying genetic (co)variances. We 37 38 tested these assumptions using three personality traits collected during 3 years on a long-term monitored breeding population of collared flycatchers, Ficedula albicollis. We partitioned the 39 40 observed phenotypic (co)variance of aggressiveness, boldness and neophobia into genetic, permanent environment and parental components, and we estimated the repeatability and 41 heritability of these traits and their between-individual correlations. All three traits were 42 repeatable between years (at least on the latent scale) but none were heritable. Permanent 43 environment effects explained 15% of the phenotypic variance in aggressiveness, and parental 44 effects explained 25% of the phenotypic variance in neophobia, in line with previous studies 45 in wild populations. The three traits showed phenotypic correlations but no between-46 individual correlations and no additive genetic covariance. Thus, our results did not support 47 the assumptions that phenotypic covariance reflects behavioural syndromes and genetic 48 covariance. We discuss the reasons for the absence of heritability and between-individual and 49

50 genetic covariance between these three personality traits in light of the possible selective51 pressures acting on this population.

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#### 54 INTRODUCTION

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Over the past two decades, personality traits, that is, repeatable between-individual 56 behavioural differences across time and contexts (Réale, Reader, Sol, McDougall, & 57 Dingemanse, 2007), have received increasing attention in animal behavioural and 58 evolutionary ecology studies (Bell, 2007; Bell, Hankison, & Laskowski, 2009; Carere & 59 Maestripieri, 2013; Dingemanse & Réale, 2005; Réale et al., 2007; Sih, Bell, & Johnson, 60 2004; Sih, Bell, Johnson, & Ziemba, 2004). Five ecologically important personality axes have 61 62 been identified to characterize the behavioural responses of individuals when interacting with their environment (activity, exploration, boldness) and with others (aggressiveness, 63 64 sociability; Réale et al., 2007). Personality traits have been shown to depend on ecological parameters (e.g. Réale et al., 2007; Sih, Cote, Evans, Fogarty, & Pruitt, 2012), to be heritable 65 (e.g. van Oers, de Jong, van Noordwijk, & Drent, 2005; van Oers & Sinn, 2013), to be linked 66 to life history traits or fitness (Dingemanse, Both, Drent, & Tinbergen, 2004; Dingemanse & 67 Réale, 2013; Duckworth & Kruuk, 2009; Reale et al., 2010; Schuett, Tregenza, & Dall, 2010; 68 Smith & Blumstein, 2008) and often to correlate with each other at the individual level, 69 forming so-called behavioural syndromes (Garamszegi, Markó, & Herczeg, 2012; Andrew 70 Sih, Bell, & Johnson, 2004; van Oers & Sinn, 2013). Such correlations may result from a 71 functional integration of personality traits favoured by selection when interactions between 72 73 these traits increase individuals' fitness in given environmental conditions (e.g. Dingemanse et al., 2007). Altogether, these various results reveal the crucial role that personality traits may 74

play in shaping evolutionary processes in wild populations (Dingemanse et al., 2004;
Duckworth & Badyaev, 2007; Karlsson Green, Eroukhmanoff, Harris, Pettersson, &
Svensson, 2016; Niemelä, Lattenkamp, & Dingemanse, 2015).

Understanding the evolution of personality traits and their associations in behavioural 78 syndromes requires understanding the mechanisms underlying these between-individual 79 differences in behaviour and their interactions, including their genetic basis. Phenotypic 80 correlations between personality traits result from the addition of between-individual 81 correlations (defining behavioural syndromes per se, whether genetic or nongenetic) and 82 within-individual (or residual) correlations (Brommer, 2013; Dingemanse & Réale, 2013; 83 Dingemanse, Kazem, Réale, & Wright, 2010). Assessing the relative contribution of between-84 and within-individual correlation components in observed phenotypic correlations can be 85 crucial because between-individual correlations may constrain the independent evolution of 86 87 the traits involved and thus may have major evolutionary consequences in the wild (Sih, Bell, & Johnson, 2004). 88

89 To reliably quantify this relative contribution of between- and within-individual correlations, multiple measurements of the personality traits considered must be collected on a 90 large number of individuals. When only single measurements are available, it is often 91 assumed that a phenotypic correlation observed between personality traits reflects an 92 underlying between-individual correlation, i.e. the within-individual correlation is negligible 93 (the so-called 'individual gambit', Brommer, 2013). Furthermore, to quantify genetic variance 94 in personality traits and genetic covariance between them, the level of genetic relatedness 95 between individuals measured has to be incorporated (e.g. via pedigree information) in the 96 models. When relatedness information is unavailable, it is often assumed that the observed 97 phenotypic (co)variance reflects the underlying genetic (co)variance (the so-called 98

99 'phenotypic gambit'; Grafen, 1984; Hadfield, Nutall, Osorio, & Owens, 2007; van Oers &
100 Sinn, 2011).

These two crucial assumptions have been tested empirically in various species, and 101 recent meta-analyses including over 30 studies, among which 25 are from wild populations, 102 have confirmed their overall validity (Brommer & Class, 2017; Dochtermann, 2011; 103 Dochtermann, Schwab, & Sih, 2015). Across these studies, the sign (and to a certain extent 104 the magnitude) of the phenotypic correlations between personality traits reliably informed on 105 106 the sign (and the magnitude) of the between-individual correlations (Brommer & Class, 2017) and of the genetic correlations (Dochtermann, 2011). Furthermore, 52% of the between-107 108 individual variation in personality traits taken separately was explained by additive genetic variance (Dochtermann et al., 2015). A recent empirical study on a wild population of yellow-109 bellied marmots, Marmota flaviventris, estimated the proportion of phenotypic (co)variance 110 111 explained by genetic, permanent environment and maternal (co)variances between four different personality traits: docility, exploration, activity and sociability (Petelle, Martin, & 112 113 Blumstein, 2015). Results showed additive genetic variations, as well as maternal and permanent environment variations, in all four traits and a positive genetic correlation between 114 activity and sociability (Petelle et al., 2015). More of such integrative studies partitioning the 115 observed phenotypic (co)variance in multiple personality traits simultaneously are needed in 116 different biological models with contrasting life histories to better understand the mechanisms 117 underlying and possibly constraining the evolution of correlated personality traits. 118

In this study, we assessed the genetic basis of three personality traits, together with the relative contribution of between- and within-individual variations in, and correlations between, these traits, chosen for their potentially important effects on crucial ecological processes (here, nest site acquisition and defence against competitors and predators) in a natural population of a small territorial, short-lived, migrant passerine bird, the collared

flycatcher, Ficedula albicollis. During 3 consecutive years, we measured for several hundreds 124 of breeding pairs in the field (1) aggressiveness towards competitors (as the agonistic reaction 125 to simulated territorial intrusions by intra- and interspecific competitors), (2) boldness 126 127 towards predators (as the latency to resume nestling feeding after human disturbance) and (3) neophobia, possibly reflecting exploration (as the latency to resume nestling feeding in the 128 presence of a novel object on the nest site; following Réale et al.'s 2007 definitions). To 129 identify the mechanisms underlying the phenotypic (co)variation in these behavioural traits, 130 we assessed to what extent (1) additive genetic, parental or permanent environment effects 131 contributed to the observed phenotypic (co)variance, while accounting for fixed individual 132 (sex, age) covariates, and (2) between-individual correlations explained phenotypic 133 correlations between these three traits. Based on many previous studies on personality traits in 134 populations of passerines of similar ecology (e.g. Dingemanse, Both, Drent, van Oers, & van 135 136 Noordwijk, 2002; Drent, Oers, & Noordwijk, 2003; Duckworth & Badyaev, 2007; Garamszegi et al., 2015; Garamszegi, Rosivall, et al., 2012; van Oers, Drent, de Goede, & van 137 138 Noordwijk, 2004), we expected heritable differences to partly explain variation in 139 aggressiveness, boldness and neophobia in our study population. Furthermore, high aggressiveness, high boldness and low neophobia may allow individuals to efficiently secure 140 and defend a breeding site when they are unfamiliar with the environment (e.g. for dispersers: 141 Cote, Clobert, Brodin, Fogarty, & Sih, 2010; Duckworth & Kruuk, 2009; Korsten, van 142 Overveld, Adriaensen, & Matthysen, 2013). Thus we predicted a functional integration and 143 (possibly genetically based) between-individual correlations between these traits. In another 144 population of collared flycatchers, male aggressiveness and boldness, but not neophobia, were 145 found to be phenotypically correlated in some years (Garamszegi et al., 2015; Garamszegi, 146 Eens, & Török, 2009). This population and ours, however, differ greatly in both demographic 147 functioning (e.g. male age structure: Hegyi, Rosivall, & Török, 2006) and selective pressures 148

149	(e.g. nest predation: Doligez & Clobert, 2003; sexual selection: Qvarnström, 1997; see also
150	Rosivall, Török, Hasselquist, & Bensch, 2004), which may affect the fitness consequences of
151	interactions between personality traits, and thus their potential functional integration.

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#### 154 METHODS

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#### 156 Study species and population monitoring

Collared flycatchers are migratory cavity nesters and readily breed in artificial nestboxes, 157 158 providing easy access to parents' identity and breeding data. Between 2011 and 2013, we conducted the behavioural tests (see below) on 1131 pairs breeding in nestboxes spread over 159 14 to 22 forest patches in our study population located on the island of Gotland (Sweden, 160 161 Baltic Sea). Each year since 1980, nests in boxes have been monitored at least weekly from late April until early July, allowing us to record major breeding variables (laying and hatching 162 163 dates; clutch size; nestling number, condition and fledging success). Breeding pairs were captured, identified and ringed if previously unringed; females were caught during incubation 164 and males while feeding nestlings. Nestlings were ringed between day 8 and day 13 after 165 hatching; fledging typically occurs 16 days after hatching. Adult and nestling identification 166 every year combined with a high return rate of both adults (approximately 40%) and juveniles 167 (approximately 10%) for such a short-lived passerine bird (Gustafsson, 1989) allowed us to 168 establish a high-quality social pedigree of the population which has previously been used in 169 170 several quantitative genetic studies (e.g. Evans & Gustafsson, 2017; Merilä & Sheldon, 2000; Sheldon, Kruuk, & Merilä, 2003; Appendix Table A1). In this population, approximately 15% 171 of all nestlings are extrapair (Sheldon & Ellegren, 1999), a percentage considered low enough 172 for quantitative genetic models to provide valid (i.e. only slightly underestimated) heritability 173

estimates from the social pedigree (Charmantier & Réale, 2005; Firth, Hadfield, Santure, Slate, & Sheldon, 2015), even though no information is available yet on how extrapair paternity may affect genetic covariances. The clear sexual dimorphism in plumage coloration in this species allowed an easy discrimination of adult males (black and white plumage with a white forehead patch) from females (brown plumage; Svensson, 1992), even from several metres away during behavioural tests. Adults could also be aged by plumage criteria (yearlings versus older adults; Svensson, 1992).

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#### 182 Aggressiveness score

We measured the level of aggressiveness of breeding flycatchers soon after settlement, during 183 nest building or early laying, that is, when the risk of losing a nestbox to a competitor is 184 highest in this single-clutch species. During the breeding season, collared flycatchers compete 185 186 for nest sites with conspecifics but also with great tits, Parus major, the second most abundant species breeding in nestboxes in the study area (Gustafsson, 1987). Aggressiveness towards 187 188 conspecific intruders was shown to decrease after the start of incubation (Král & Bícík, 1989) 189 even though aggressiveness towards great tit intruders remained high throughout the breeding cycle (Král & Bícík, 1992). To elicit an aggressive response from a focal flycatcher pair, we 190 simulated the intrusion of competitors at the nest of the pair by attaching to its nestbox clay 191 decoys mimicking either a flycatcher pair or a single (male) great tit. We used a pair (one 192 male and one female) for flycatcher decoys to elicit and measure an aggressive response by 193 both pair members, that is, to avoid a sex-specific response towards this intraspecific 194 stimulus, while the response to the interspecific stimulus (great tit decoy) was not expected to 195 differ depending on the sex of the decoy. In addition, we simultaneously broadcast male songs 196 197 corresponding to the decoy(s) species with a loudspeaker placed just under the nestbox. To avoid pseudoreplication, we randomly used one of eight different sets of decoys and one of 198

five different song tracks per species for each test. After attaching the decoys to the nestbox 199 and the loudspeaker under the box, the observer sat under a camouflage net approximately 8-200 10 m away from the nestbox and recorded the following behaviours for each pair member: (1) 201 202 movements between perches and perching position (within 2 m, between 2 and 5 m or between 5 and 10 m away from the nestbox), (2) agonistic behaviours towards a decoy 203 (attacks and stationary flights in front of the decoy) and (3) chases towards living birds 204 attracted by the stimulus. A behavioural test started with an observation period of 15 min but 205 206 we lengthened the test by up to 5 min when an individual arrived between 10 and 15 min after the start of the test, and up to 5 additional min if its partner arrived during this extra time, so 207 that we could observe the behavioural response of each pair member for at least 5 min. Each 208 test thus lasted between 15 and 25 min. If an individual was observed during less than 5 min 209 before the end of the test, it was discarded from the analyses and these observations were 210 211 therefore not used later on.

Aggressiveness level was measured as the sum of the number of movements within 2 212 213 m of the nestbox, attacks, stationary flights and chases. We included this latter behaviour because chasing a live intruder may have prevented the focal flycatcher from interacting with 214 the decoy, while reflecting an aggressive territory defence response. The number of each type 215 of behavioural response (movements, agonistic behaviours and chases) was standardized by 216 the time interval between the first observation of the individual and the end of the test, 217 rescaled to 15 min. Using alternative scores did not qualitatively change the results (see 218 Appendix and Table A2). We conducted aggressiveness tests two to four times per focal pair 219 220 over a 5-day interval, with at most one test per day and tests on 2 days in a row. The stimulus used (intra- / interspecific decoys) was alternated between tests after a random assignment for 221 the first test. An aggressiveness score was computed for each individual for each test. We 222 obtained aggressiveness responses (for more than 5 min at least once per year) for 1974 223

individuals (including unidentified ones; 961 females and 1014 males in 1046 nests). Among
those, 825 females and 667 males were later captured and identified, and thus used for
heritability estimation. In 601 breeding pairs both partners were identified and responded to
the tests and in 273 only one partner responded. We obtained repeated estimates for 502 and
445 identified females and males, respectively.

229

#### 230 Boldness and neophobia scores

During nestling rearing, we estimated (1) boldness level by measuring the individual's 231 reaction towards the presence of a human observer near the nestbox and (2) neophobia level 232 by measuring the reaction towards the presence of a novel object on the nestbox (i.e. in a 233 familiar environment), following the definitions from Réale et al. (2007). As advised in 234 Greenberg and Mettke-Hofmann (2001), we measured our behavioural responses as the 235 236 latency to perform a highly motivating action (here, feeding their nestlings) after disturbance. We conducted one combined boldness/neophobia test per breeding pair when the nestlings 237 238 were 5 days old, that is, at the beginning of the period of highest provisioning by parents (and 239 before we caught the parents to avoid any behavioural interference). A test consisted of two consecutive periods of approximately 1 h each: the behaviour of the parents was recorded first 240 without any change in the surroundings of the nestbox, that is, without the novel object, and 241 second with a novel object (here a coloured figurine approximately 7 cm high) attached near 242 the entrance hole of the nestbox. Both periods were video-recorded from a distance (6–8 m). 243 At the beginning of each period, the observer checked the camouflaged video recorder, 244 walked to the nestbox, opened it to check nestling satiety, closed it, and then left the area. The 245 test was abandoned if the nestlings were very hungry to avoid them starving if the parents 246 247 were too disturbed by the test.

We estimated boldness score using the latency to enter the nestbox after the departure 248 of the observer in the first period (i.e. without the novel object). Reaction to disturbance by 249 humans has previously been used in boldness tests in this species (e.g. Garamszegi et al., 250 2009). To ease interpretation (i.e. increasing values of boldness score corresponding to 251 increasing level of boldness), we transformed the latency to enter the nestbox such that the 252 boldness score of an individual was the maximum latency observed in the entire data set 253 minus the latency for this individual. We estimated neophobia score based on the latency to 254 255 enter the box after the departure of the observer in the second period (i.e. in the presence of the novel object). For both boldness and neophobia scores, individuals that did not enter the 256 nestbox during the first period of the test were not used in the analysis (187 of 1251 257 observations, i.e. 15%). Individuals that entered the nestbox during the first but not the second 258 part of the test (411 of 1064 observations, i.e. 39%) were considered as the most neophobic 259 260 ones but could not be assigned a latency. To include them in the analyses, we discretized the latency to enter the nestbox in the second period into four categories based on its quartiles, 261 262 assigning values from 1 to 4 for increasing latencies, and adding a fifth category including 263 individuals that did not enter in the second part of the test. Using alternative scores for boldness and neophobia did not qualitatively change the results (see Appendix and Tables A2 264 and A3). We obtained boldness and neophobia estimates for 849 identified individuals (472 265 females and 378 males). Over the 3 years, we assessed 403 unique breeding pairs where both 266 identified partners responded to the tests and 185 pairs where only one partner responded. We 267 obtained repeated boldness and neophobia scores (i.e. several years in a row) for 66 females 268 and 65 males. 269

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#### 271 Repeatability and heritability of aggressiveness, boldness and neophobia

We estimated the repeatability of aggressiveness, boldness and neophobia scores as well as 272 273 their heritability by fitting three separate univariate animal models. The models included the following random effects: additive genetic effect (associated with the pedigree), individual 274 identity for the repeated measures per individual (permanent environment effect), maternal 275 and paternal identities, forest patch, observer identity (the person observing and reporting the 276 behaviours onsite for the aggressiveness tests and the person extracting latencies from the 277 video recording for the boldness and neophobia tests). The models of the aggressiveness score 278 also included the broadcast song track and decoy set identifiers. In addition, the models 279 included the following fixed effects, to control for potential confounding factors: sex, age 280 281 (two levels: yearling versus older) and their interaction, as well as year (three levels: 2011, 2012, 2013). The aggressiveness model also included stimulus type (two levels: flycatcher 282 versus great tit decoys), the order of the test within a year (continuous variable: first to 283 fourth), the presence of the partner during the test (binary variable: yes/no) and the presence 284 of other live flycatchers or great tits (binary variable: yes/no). The boldness and neophobia 285 models included the number of ringed nestlings as a proxy of the motivation to enter the 286 nestbox to feed nestlings. The neophobia models included the boldness score to control for the 287 effect of the human disturbance at the beginning of the period with the novel object. All 288 289 continuous fixed terms were centred and standardized prior to analysis to allow comparisons between effects. 290

Repeatabilities (*R*) were estimated as the ratio of the sum of the additive genetic ( $V_A$ ), permanent environment ( $V_{PE}$ ), maternal and paternal identities variances ( $V_M$  and  $V_F$ , respectively) over the total phenotypic variance ( $V_P$ , sum of all variances; Falconer & Mackay, 1996; with possibly an additional term accounting for the distribution variance, Nakagawa & Schielzeth, 2010). Narrow-sense heritabilities ( $h^2$ ) were estimated as the ratio of the additive variance  $V_A$  over the phenotypic variance  $V_P$ . The presence of fixed effects in models did not result in over- or underestimating repeatability and heritability estimates (as warned against in Nakagawa & Schielzeth, 2010), because estimates were similar when obtained from models with only the intercept (presented in the main text) and from models with the previously described fixed effects (see Appendix Tables A1, A2, A3; except for a slight overestimation of the boldness and neophobia repeatabilities on the latent scale). For aggressiveness, we also estimated within-year repeatability by replacing the additive genetic and permanent environment effects by a unique identifier per individual per breeding season.

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#### 305 Correlations between aggressiveness, boldness and neophobia scores

We estimated the between- and within-individual correlations between aggressiveness, 306 boldness and neophobia scores by fitting a trivariate mixed-effects model (Dingemanse & 307 Dochtermann, 2013). For this model, aggressiveness was averaged over all aggressiveness 308 309 scores of an individual in a given year (i.e. over up to four estimates). We chose this averaging approach because aggressiveness score (1) differed depending on decoy species, 310 311 chosen at random for the first test, and (2) decreased with the order of the test (i.e. due to 312 habituation). Averaging all aggressiveness scores of an individual in a given year allowed us to control for these differences and decrease the effect of varying environmental 313 (meteorological) conditions between tests. In this trivariate model, we included sex and year 314 as fixed effects and ring number as a random effect. The phenotypic correlation between two 315 traits A and B,  $r_{P_A,P_B}$ , and the between-individual correlation between the traits A and B, 316  $r_{\text{ind}_{A},\text{ind}_{B}}$ , were estimated as follows (Snijders & Bosker 1999): 317

318 
$$r_{P_A,P_B} = \frac{Cov_{ind_A,ind_B} + Cov_{\varepsilon_A,\varepsilon_B}}{\sqrt{V_{P_A} \times V_{P_B}}}$$
(1)

319 
$$r_{\text{ind}_A, \text{ind}_B} = \frac{\text{Cov}_{\text{ind}_A, \text{ind}_B}}{\sqrt{V_{\text{ind}_A} \times V_{\text{ind}_B}}}$$
 (eq. 2)

where  $Cov_{ind_A,ind_B}$  and  $Cov_{\epsilon_A,\epsilon_B}$  are the between- and the within-individual covariances 320 between traits A and B, and  $V_{P, A \text{ or } B}$  the total phenotypic variance (sum of the between- and 321 within-individual variances) associated with trait A or B. Combining aggressiveness, boldness 322 and neophobia score data for a given individual in a given year, we obtained 841 observations 323 324 for which estimates for all three traits were available, and 152 observations for which only boldness and neophobia estimates were available. Among these 841 observations, 49 females 325 and 52 males were repeatedly assessed over several years for all three traits, providing 100 326 327 and 111 repeated observations, respectively.

To estimate the additive genetic correlations between traits, we fitted a model with the same fixed effects but with the additive genetic effect instead of the individual ring as a random factor. Fitting both additive genetic and individual (permanent environment) effects together in a single model indeed led to convergence failures. We thus fitted only one effect at a time.

333

#### 334 Implementation of Bayesian models

335 All statistical analyses were performed within the Bayesian framework in R v.3.3.2 (R Core Team, 2016). Both univariate and trivariate models were fitted using the function 336 MCMCglmm ('MCMCglmm' R package, Hadfield, 2010). The pedigree was prepared using 337 the function fixPedigree ('pedantics' R package, Morrissey & Wilson, 2010) and pruned 338 using the function prunePed ('nadiv' R package, Wolak, 2012; see Appendix Table A1 for a 339 description of the pedigree). Aggressiveness and boldness scores were fitted with a Poisson 340 family (logit link), and neophobia scores with a Threshold family with the residual variance 341 fixed to 10 (instead of the usual value of 1, to improve the mixing of the chains for low 342 variances, which were expected from preliminary analyses; Hadfield, 2016). We used wide 343 normally distributed priors for fixed effects (large variance  $V=10^8$ ; Hadfield, 2016) and 344

parameter expanded  $\chi^2$  distributed priors with 1 degree of freedom for random effects. For the 345 univariate models, we adjusted the number of iterations, burn-in and thinning interval for each 346 model so as to obtain an effective sample size over 1500 (see Appendix) and autocorrelations 347 of posterior samples below 0.1 in all cases. For the trivariate models, we used 4 x  $10^6$ 348 iterations, a burn-in of 10<sup>5</sup> and a thinning interval of 2000 to reach the same criteria. We 349 visually assessed the convergence of each MCMC chain and compared three chains per model 350 using the Gelman & Rubin approach (gelman.diag and gelman.plot functions, 'coda' R 351 package, Plummer et al., 2016). Following Nakagawa and Schielzeth (2010), we retained in 352 our univariate models data from individuals tested only once. For all three traits, estimates are 353 354 presented as posterior modes with the associated 95% credible intervals (CI). Variance, repeatability and heritability estimates are presented on the latent scale ( $R_{\text{latent}}$ , following 355 Nakagawa & Schielzeth, 2010; h<sup>2</sup><sub>latent</sub> following de Villemereuil, Schielzeth, Nakagawa, & 356 357 Morrissey, 2016). We also provide repeatability and heritability estimates on the observed scale ( $R_{obs}$  using the QGicc function,  $h^2_{obs}$  using the QGparams function, from the 'QGglmm' 358 R package; de Villemereuil, 2018; de Villemereuil, Schielzeth, Nakagawa, & Morrissey, 359 2015). It is not possible here to discuss all results on the observed scale, even though it is the 360 scale of the realized behaviour, that is, where natural selection can act (see de Villemereuil et 361 al., 2016). Indeed, the estimation of  $R_{obs}$  for ordinal traits (here the neophobia score) is 362 complex and is not currently implemented in QGicc. To allow us to compare estimates 363 between traits on the observed scale, we computed repeatabilities for neophobia using two 364 alternative scores with nonordinal distributions, namely (1) the latency to enter the nestbox in 365 the presence of the novel object for individuals that entered the nestbox (N=653, fitted with a 366 Poisson distribution), and (2) whether the individual entered the nestbox in the presence of the 367 novel object (N=1064, binary variable fitted with the threshold family). Moreover, the 368 estimation of  $h_{obs}^2$  for ordinal traits provides one heritability estimate per level, which in the 369

case of an artificial categorization as here is not biologically relevant (de Villemereuil, 2018).
As the heritability estimates on the observed scale were fairly similar between neophobia
levels, we compared the range of heritability estimates found for neophobia with the
heritability estimates found on the observed scale for aggressiveness and boldness.
Correlations are provided on the latent scale.

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#### 376 Ethical note

Permission for catching and ringing adults (here 838 yearlings, 1074 older birds) and 377 nestlings (here 9750) with individually numbered aluminium rings was granted every year by 378 the Ringing Centre from the Museum of Natural History in Stockholm (licence nb. 471: M015 379 to B.D.). Adults were caught in the nest, either directly (females during incubation) or using 380 swinging-door traps (both parents during nestling rearing). Traps were set for at most 30 - 60 381 382 min depending on nestling age (30 min when nestlings were 5 days old or younger), to avoid nestling starvation if parents did not resume feeding during the catching period; traps were 383 384 checked every 5-10 min, and removed as soon as the adults had been caught. Catching sessions started after 0600 hours to let birds feed and provision nestlings undisturbed for at 385 least 2 h after the night period (sunrise occurs at approximately 0400 hours during spring). 386 Adults were handled for 5 - 10 min and released straight after manipulation or (when catching 387 both parents during nestling feeding) kept until the partner was captured (up to 40 min 388 maximum). For nestling ringing, whole broods were taken directly from the nest and ringed 389 just beside the nestbox (for approximately 10 min); nestlings that were not handled were kept 390 391 warm using small heating packs. During the aggressiveness tests, we minimized disturbance by approaching the nestbox as quietly as possible and hiding below a camouflage net. 392 Conversely, for the combined boldness/neophobia test, which aimed at measuring (or 393 controlling for) the reaction towards human presence, we on purpose approached the nestbox 394

conspicuously. During the boldness/neophobia test, nestling satiety was checked at the beginning and in between the two parts of the test, and the test was aborted if nestlings were begging too strongly, to avoid any harmful effect of temporarily decreased provisioning by parents. All the manipulations were performed in accordance with the Swedish legislation applying at the time.

- 400
- 401
- 402 **RESULTS**
- 403

#### 404 Univariate models

The level of repeatability for aggressiveness was 0.18 on the latent scale and 0.03 on the 405 observed scale (Table 1). In addition, aggressiveness score was repeatable within years 406 407  $(R_{\text{latent}}=0.22, 95\% \text{ CI} = [0.18; 0.26]; R_{\text{obs}}=0.04, 95\% \text{ CI} = [0.03; 0.06])$  and between years (when averaging the aggressiveness score of 1 year;  $R_{\text{latent}}=0.26$ , 95% CI = [0.11; 0.38]; 408 409  $R_{\rm obs}$ =0.11, 95% CI = [0.04; 0.20]). The level of repeatability for boldness was 0.11 on the latent scale and 0.10 on the observed scale (Table 1). Neophobia was slightly more repeatable 410 that the other traits on the latent scale (Table 1) as well as on the observed scale when 411 estimated from alternative nonordinal measures ( $R_{\text{latent}}=0.25$ , 95% CI = [0.06; 0.35] and 412  $R_{obs}=0.14$ , 95%CI = [0.04; 0.23] for the latency to enter the nestbox in the presence of the 413 novel object;  $R_{\text{latent}}=0.44$ , 95% CI = [0.22; 0.67] and  $R_{\text{obs}}=0.29$ , 95%CI = [0.12; 0.44] for 414 whether the individual entered the nestbox during the test or not). None of the three 415 behavioural scores, however, were heritable (all 95% CI for  $V_A$  and  $h^2$  values included zero; 416 Table 1). Permanent environment effects explained 15% of the phenotypic variance in 417 aggressiveness score (95% CI of  $V_{PE} = [0.28; 0.80]$ ; Appendix Table A2) and parental 418 identities (i.e. maternal and paternal identities summed) explained 25% of the phenotypic 419

variance in neophobia score (95% CI of  $V_{\rm M} + V_{\rm F} = [0.38; 8.46]$ , even though the lower limit of 420 the 95% CI for each parent identity separately was 0: [0.00; 6.49] for  $V_{\rm M}$  and [0.00; 5.75] for 421  $V_{\rm F}$ ). When excluding the maternal (paternal) identity from the model, the paternal (maternal) 422 identity explained 19% (18%) of the phenotypic variance. Fitting the neophobia model 423 without the maternal and paternal identities did not change the heritability estimate, revealing 424 that these effects were not confounded with the additive genetic effect (results not detailed). 425 Observer identity explained 7% of the phenotypic variance for aggressiveness (95% CI of 426  $V_{\text{Observer}} = [0.14; 0.55]$ ; paternal identity explained 11% of the phenotypic variance for 427 aggressiveness but only when measuring aggressiveness as the first axis of a principal 428 component analysis (see Appendix). All other variances were low (less than 4 % of the 429 phenotypic variance) or not different from zero (Appendix Tables A1, A2 and A3). 430

Males were more aggressive than females, especially among yearlings (interaction 431 432 sex\*age, with yearling males as reference: 95% CI = [0.44, 0.81]; Appendix Fig. A1a). In addition, males were slightly shyer (longer latency to enter in the absence of a novel object) 433 434 and less neophobic (shorter latency in the presence of a novel object) than females (with female as a reference 95% CI = [-0.12; -0.02] and [-2.78; -1.07] respectively; Appendix Fig. 435 A1b), and this did not depend on age (see Appendix Tables A2 and A3 for sex\*age 436 interactions). In addition, individuals with larger broods were bolder (95% CI = [0.002; 0.06]; 437 Table A3) and less neophobic (95% CI = [-1.03; -0.37]; Table A4). Regarding environmental 438 effects, aggressiveness, boldness and neophobia scores depended on the year: individuals 439 were less aggressive and less neophobic in 2011 than 2012 (Appendix Tables A2 and A4), 440 and shyer in 2011 than 2013 (Appendix Table A3). Finally, individuals were more aggressive 441 in the presence of their partner or neighbouring tits attracted by the stimulus, and during the 442 first tests of the sequence (Appendix Table A2). 443

#### 445 *Trivariate model*

Aggressiveness and neophobia scores were phenotypically correlated: more aggressive 446 individuals were less neophobic (Table 2, Appendix Fig. A2a). Boldness and neophobia 447 448 scores were also phenotypically correlated: bolder individuals were less neophobic (Table 2, Appendix Fig. A2b). The corresponding within-individual correlations were negative (Table 449 2). However, there was no phenotypic correlation between boldness and aggressiveness 450 (Table 2; see Appendix Table A5 for the full model output) and none of the between-451 individual correlations differed from zero (Table 2). When accounting for an additive genetic 452 effect, none of the additive genetic covariances differed from zero (Table 2; see Table A6 for 453 the full model output). 454

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457 DISCUSSION
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459 In this study, we tested whether three personality traits (aggressiveness, boldness and 460 neophobia) had a genetic basis in a wild population of collared flycatchers and formed (genetically based) behavioural syndromes during breeding, to shed light on constraints in 461 their possible evolution. None of the three traits were heritable and their repeatability 462 estimates were low (0.11–0.39 on the latent scale for all traits; 0.03–0.10 on the observed 463 scale for aggressiveness and boldness; 0.14 and 0.29 on the observed scale for nonordinal 464 measures of neophobia) compared to average estimates previously found for behavioural traits 465 in two meta-analyses (0.37 in Bell et al., 2009; 0.41 in Holtmann et al., 2017), suggesting 466 strong phenotypic plasticity in these traits. The repeatability originated mainly from 467 permanent environment effects for aggressiveness and from parental effects for neophobia. In 468 addition, we found that the three traits showed phenotypic covariance but no between-469

individual covariance and no additive genetic covariance. The absence of behavioural
syndromes among these personality traits may be due either to a lack of statistical power to
detect between-individual covariances or to an absence of functional integration of these traits
at the individual level and no genetic correlation at the population level.

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### 475 Factors at the origin of behavioural trait repeatability

Our levels of repeatability, estimated both within and between years for aggressiveness score 476 477 and between years for boldness and neophobia scores, were lower than usually reported for such behaviours: around 0.50 for aggressiveness and exploration, and around 0.40 for 478 antipredator behaviours (Bell et al., 2009). Interestingly, the repeatability level estimated here 479 for aggressiveness score was similar within and between years, contrary to the usual decrease 480 observed when the time interval between recordings increases (Bell et al., 2009; Chervet, 481 482 Zöttl, Schürch, Taborsky, & Heg, 2011; Dingemanse et al., 2012; Garamszegi et al., 2015; Holtmann et al., 2017; Wuerz & Krüger, 2015; but see David, Auclair, & Cézilly, 2012 for 483 484 differences between traits). Overall, our lower levels of repeatability, especially on the observed scale, suggest higher plasticity, both within and between years compared to studies 485 on other species. 486

The observed repeatability in aggressiveness score resulted mostly from permanent 487 environment effects, which explained 15% of the phenotypic variance and 72% of the 488 repeatability in aggressiveness score. Here, because we controlled for the identity of the 489 parents, permanent environment effects could be linked to differences in individual condition 490 or experience. Some measures of personality traits have indeed been found to depend on 491 individual condition or experience (reviewed in Stamps & Groothuis, 2010). In our 492 population, condition and experience also affect breeding habitat choice depending on social 493 cues (e.g. Doligez, Danchin, Clobert, & Gustafsson, 1999; Doligez, Pärt, Danchin, Clobert, & 494

Gustafsson, 2004; Kivelä et al., 2014), which could shape individuals' response to the risk of competition for nest sites. Permanent environment effects may also include a dominance effect (Kruuk & Hadfield, 2007; Wilson et al., 2010), which could not be directly modelled here because full- and half-sib links were too rare in our pruned pedigree for running such complex models (Wilson et al., 2010).

In turn, the observed repeatability in neophobia resulted mostly from parental 500 identities, which accounted for 25% of the phenotypic variance and 55% of the between-501 502 individual variance. Both pre- and postnatal parental effects have been found to affect exploration and neophobia behavioural responses later in life (e.g. nestling provisioning and 503 exploration in birds, Carere, Drent, Koolhaas, & Groothuis, 2005; maternal hormones early in 504 life and neophobia, Spencer & Verhulst, 2007; see the review in Groothuis & Maestripieri, 505 2013). However, parental identities did not explain between-individual differences in 506 507 aggressiveness and boldness here, contrary to previous findings (e.g. Eising, Muller, & Groothuis, 2006; reviewed in Groothuis & Maestripieri, 2013). These behavioural responses 508 509 may be more dependent on individual or local environmental conditions, in particular 510 individual competitive ability and neighbour/predator presence or density, at the time of the test(s). To better understand how parental effects shape behavioural responses later in life in 511 our study population, further experiments (e.g. nestling cross-fostering) would be necessary. 512

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#### 514 Personality traits with no genetic basis

We found no genetic basis for our three personality traits. A meta-analysis on personality traits in wild animal populations estimated an average heritability level of 0.28 for aggressiveness, 0.31 for boldness and 0.58 for exploration-avoidance (including estimates from novel environment and novel object tests; van Oers & Sinn, 2013). The absence of heritability for our personality traits here was not due to a lack of statistical power to detect

significant additive genetic variance based on our sample and social pedigree, because based 520 on the same sample with the same pedigree, we obtained positive heritability estimates for 521 tarsus and wing length ( $h^2=0.59$ , 95% CI = [0.44; 0.69] for tarsus and  $h^2=0.30$ , 95% CI = 522 [0.11; 0.50] for wing length, while accounting for maternal and permanent environment 523 effects) which are consistent with previous estimates in this population ( $h^2 = 0.53$  and 0.51 for 524 tarsus and wing length, respectively, in Merilä & Gustafsson, 1993). The absence of 525 heritability in our personality traits was therefore likely to be the result of very low additive 526 genetic variance combined with large environmental variance as illustrated, for instance, by 527 between-year differences in behavioural scores, which reflected large variations in 528 environmental conditions between the 3 years of our study (see Morinay, Forsman, Kivelä, 529 Gustafsson, & Doligez, 2018 for differences between 2012 and 2013). Large environmental 530 variance could originate from individuals being tested in different environments (including 531 532 the social context) in different years, because between-year fidelity to the nestbox and/or partner is very low in this population (approximately 6.7% of 240 individuals bred in the 533 534 same nestbox several years and 1.0% of 214 identified pairs were faithful over several years). This, however, limited the risk of pseudoreplication (Niemelä & Dingemanse, 2017). 535

Because we measured personality traits at the nest during breeding, the reaction of the 536 partner may have affected the reaction of the focal bird during a behavioural test, as found 537 here with a higher aggressiveness score when the partner was present. To account for this 538 effect, we could have included the partner's identity and genetic background (i.e. pedigree) as 539 random effects in our models (see Morinay et al., 2018 for an example in the same 540 population; and Wolf, Brodie III, Cheverud, Moore, & Wade, 1998 for so-called indirect 541 genetic effects). However, this could have led to pseudoreplication, because the behavioural 542 score of the partner itself was most of the time also analysed in this dataset (e.g. around 74% 543 of the females and 90% of the males had their partner tested). To keep exploring a response at 544

the individual level (rather than combining behavioural scores at the pair level), a solution 545 could be to fit a bivariate model of the two partners' responses and include both their 546 pedigree, permanent environment and parental effects in the model. This would, however, 547 require a larger data set than used here to reach sufficient statistical power to detect such 548 effects with such complex models. Furthermore, because the focal bird chooses at least partly 549 its partner (like its nest site), partner's effects can also be expected to be at least partly 550 included in the individual's genetic and permanent environment effects. Disentangling such 551 complex effects may require a more balanced sample of faithful and divorced pairs breeding 552 in the same and different sites over several years than observed in our population. 553

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#### 555 No personality syndrome?

Phenotypic correlations were observed between our personality traits, even though they did 556 557 not constitute behavioural syndromes (i.e. there was no between-individual correlations): less neophobic individuals were more aggressive and bolder. This was in line with previous 558 559 studies reporting bolder individuals to be more explorative in a novel environment (or less neophobic in a novel object test; e.g. Garamszegi et al., 2009; van Oers, De Jong, Drent, & 560 van Noordwijk, 2004). Conversely, the absence of correlation between aggressiveness and 561 boldness partly contrasts with previous results reporting more aggressive individuals to be 562 bolder, as part of the proactive-reactive axis, in different species (Koolhaas et al., 1999; Sih, 563 Bell, Johnson, et al., 2004) including the collared flycatcher (Garamszegi et al., 2015). 564

The observed phenotypic correlations resulted solely from correlated changes in behaviours between measurements for the same individuals, that is, within-individual correlations. Within-individual correlations could be due to micro-environmental effects (e.g. nestbox environment), to individual effects (e.g. long-term between-year plasticity but shortterm within-year behavioural constraints, for instance due to experience) or to correlated

measurement errors (Dingemanse & Dochtermann, 2013). Error correlation, however, is more 570 likely to occur between boldness and neophobia scores, which were extracted from the same 571 test and might both be correlated with feeding rate (with birds investing more in nestling 572 573 provisioning returning more rapidly to their nest in both situations), than between aggressiveness and neophobia scores, which were measured several weeks apart by different 574 persons in different tests. To tease these sources of within-individual correlations apart, 575 aggressiveness, boldness and neophobia scores need to be estimated several times during the 576 577 same breeding season and possibly the same phase(s) of the reproductive cycle. The limited number of individuals measured several times here (211 observations of 101 individuals), 578 579 however, is likely to be why we did not detect between-individual covariance (Dingemanse & Dochtermann, 2013 recommended sample sizes of at least 200 individuals tested twice; see 580 also Garamszegi & Herczeg, 2012). Indeed, based on the same sample, we were not able to 581 582 obtain positive genetic or between-individual covariances between tarsus and wing length, two morphological traits previously reported as genetically correlated in the same population 583 584 (Merilä & Gustafsson, 1993).

Even though our limited statistical power does not allow us to conclude the absence of 585 behavioural syndromes, this absence, if true, would suggest that selective pressures did not 586 yield or maintain a functional integration between the personality traits investigated here. A 587 true absence of behavioural syndromes among the traits we studied could be explained by 588 specific breeding conditions in our population, possibly altering the selective regimes 589 compared to other populations or species. In our population, the high availability of high-590 591 quality nest sites (i.e. nestboxes, provided in excess since the early 1980s) may have released joint selective pressures on exploration to find suitable nest sites and aggressiveness to 592 acquire and defend this resource against dominant competitors (in particular tit species) in a 593 natural context. In turn, providing nestboxes probably increased local breeding densities and 594

thereby competition for food resources during the nestling period, especially in a highly synchronous species such as the collared flycatcher. Furthermore, our population is subjected to very low nest predation rates, due to the absence of mustelid species on Gotland (Doligez & Clobert, 2003), which may have released selective pressures on boldness through the decrease in the need to defend the brood. Overall, these specific breeding conditions may have strongly modified the selective regime for personality traits and for a functional integration between them if they are costly.

In conclusion, we showed that aggressiveness, boldness and neophobia are repeatable 602 but not heritable traits and do not seem to form behavioural syndromes in our population of 603 collared flycatchers since only phenotypic correlations were observed between neophobia and 604 the other two traits. Our study thus brings insights on the evolutionary potential of these 605 personality traits alone and in interaction with each other during breeding in a wild population 606 607 experiencing particular breeding conditions (low competition for nest sites, low nest predation rate). To understand the absence of individual covariance between, and heritable variations in, 608 609 personality traits in our population, a first step would be to investigate the fitness benefits (i.e. 610 reproductive success and survival) associated with each trait and their interactions. Plasticity in the associations between personality traits should be selected for if the fitness costs and 611 benefits of expressing each trait relative to the others depend on the environmental (including 612 social) context (e.g. competition level or predation risk), which remains to be explored for 613 example by experimentally manipulating these environmental conditions. 614

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#### 616 **Conflict of interest**

617 We declare no conflict of interest.

618

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899

#### 900 Appendix

### 901 <u>Aggressiveness score</u>

We estimated aggressiveness score using the following alternative measures and modelled 902 them using the parameters given in parentheses. (1) Number of aggressive behaviours 903 (movements < 2 m from the nestbox, stationary flights and attacks towards the decoys, chases 904 of live birds) standardized per 15 min (Poisson family; number of iterations =  $10^6$ ; burn-in = 905  $10^4$ ; thinning interval = 400). (2) First axis of the PCA presented below (Table A7; Gaussian 906 family; number of iterations =  $10^6$ ; burn-in =  $10^4$ ; thinning interval = 400). (3) Discrete score 907 (threshold family, residual variance  $V_{\rm R} = 10$ ; number of iterations = 13 x 10<sup>5</sup>; burn-in = 8 x 908  $10^4$ ; thinning interval = 700). This score was based on the distinction between activity 909 910 (number of movements/min, including stationary flights and chases) performed far from (> 2 m) and close to (< 2 m) the nestbox, and on attacks, subdivided into six categories (Fig. A3a): 911 0: individuals that performed no movements (either far from or close to the nestbox); 1: 912 individuals that performed no attack or movements close to the nestbox and performed less 913 than 0.440 movements/min far from (> 5 m) the nestbox; 2: individuals that performed no 914 attack or movements close to the nestbox and performed more than 0.440 movements/min far 915 from (> 5 m) the nestbox; 3: individuals that performed no attacks and less than 0.282 916 movements/min close to (< 2 m) the nestbox; 4: individuals that performed no attacks and 917 between 0.282 and 0.784 movements/min close to (< 2 m) the nestbox; 5: individuals that 918 performed no attacks and above 0.784 movements/min close to (< 2 m) the nestbox; 6: 919

920 individuals that performed attacks towards decoys(s). The thresholds were chosen so as to
921 distribute individuals equally among categories for scores 1 and 2 on the one hand and scores
922 3, 4 and 5 on the other.

923

### 924 <u>Boldness score</u>

We estimated boldness score using the following alternative measures and modelled them 925 using the parameters given in parentheses. (1) Maximum latency to enter the nestbox after 926 human departure (from all individuals) minus the same latency for the focal individual on the 927 focal test (Poisson family; number of iterations =  $1 \times 10^6$ ; burn-in =  $10^4$ ; thinning interval = 928 929 500). (2) Inverse logarithmic ratio of the latency to enter the nestbox after human departure divided by the feeding rate for the remaining time after the first entrance in the nestbox. We 930 divided by the feeding rate for this alternative measurement because the average feeding 931 932 interval might have affected the latency to return to the nestbox, for example if individuals feeding more frequently entered the nestbox faster. The feeding rate after the first entrance 933 934 was estimated as the average time interval between two feeding events by the focal birds, after the first entrance; it was thus only computable for individuals that fed at least twice in 935 the period; for technical reasons, further individuals could not be used for this variable, 936 leading to a final used data set of 641 observations (Gaussian family; number of iterations = 937 25 x 10<sup>4</sup>; burn-in = 10<sup>4</sup>; thinning interval = 100). (3) Discrete score (threshold family,  $V_{\rm R}$  = 938 10; number of iterations =  $3 \times 10^6$ ; burn-in =  $2 \times 10^5$ ; thinning interval = 1000). The score was 939 based on entrance in the nestbox during the first part of the boldness-neophobia test (no novel 940 object) and latency to enter after human disturbance; individuals that did not enter were given 941 a score of 0, and individuals that entered the nestbox were given a score of 1 - 5 based on five 942 quantiles of the inverse latency to enter (Fig. A3b). 943

### 946 <u>Neophobia score</u>

We estimated neophobia score for individuals that entered the nestbox during the first period 947 of the test, using the following alternative measures and modelled them using the parameters 948 given in parentheses. (1) Discrete score based on the latency to enter the nestbox in the 949 presence of a novel object (second period of the test), discretized in four quantiles, the fifth 950 category including individuals that did not enter the nestbox at all in the presence of the novel 951 object (threshold family;  $V_{\rm R} = 10$ ; number of iterations = 15 x 10<sup>5</sup>; burn-in = 2 x 10<sup>5</sup>; thinning 952 interval = 500; Fig. A3c); (2) Binary variable separating individuals that did and did not enter 953 during the second period of the test (threshold family;  $V_{\rm R} = 10$ ; number of iterations =  $10^6$ ; 954 burn-in =  $10^5$ ; thinning interval = 500). (3) Latency to enter the nestbox in the presence of the 955 novel object, excluding the individuals that did not enter the nestbox during the second period 956 of the test (Poisson family; number of iterations=  $15 \times 10^5$ ; burn-in =  $15 \times 10^4$ ; thinning 957 interval = 500). 958

959 Results obtained for these alternative scores for the three behavioural traits are given in960 Tables A2 to A4, and the main text presents the first score in each case.

### 961 TABLES

962

963 Table 1. Between-year repeatability and heritability estimates for aggressiveness, boldness

R <sub>latent</sub>	$R_{ m obs}$	$h^2_{\text{latent}}$	$h^2_{ m obs}$
0.18 *	0.03 *	0.00	0.00
[0.15; 0.23]	[0.02; 0.04]	[0.00; 0.08]	[0.00; 0.008]
0.11 *	0.10 *	0.00	0.00
[0.01; 0.21]	[0.01; 0.19]	[0.00; 0.10]	[0.00; 0.09]
0.39 *		0.00	0.00 for all scores
[0.25, 0.54]		[0, 00, 0, 1 <b>5</b> ]	From [0.00; 0.00]
[0.23; 0.34]		[0.00; 0.13]	to [0.00; 0.09]
	0.18 * [0.15; 0.23] 0.11 * [0.01; 0.21]	0.18 *       0.03 *         [0.15; 0.23]       [0.02; 0.04]         0.11 *       0.10 *         [0.01; 0.21]       [0.01; 0.19]         0.39 *	0.18 *0.03 *0.00[0.15; 0.23][0.02; 0.04][0.00; 0.08]0.11 *0.10 *0.00[0.01; 0.21][0.01; 0.19][0.00; 0.10]0.39 *0.00

964 and neophobia scores

965

966 Repeatabilities and heritabilities (posterior modes and 95% credible intervals) are given on the latent scale ( $R_{latent}$ ,  $h^{2}_{latent}$ ) and on the observed scale ( $R_{obs}$  obtained with the QGicc function 967 and  $h^2_{obs}$  with the QG params function from 'QG glmm' R package; de Villemereuil et al. 968 969 2016; de Villemereuil 2018). Asterisks indicate estimates whose 95% CI do not encompass zero. For aggressiveness, estimates are given using all scores. For neophobia, we provide a 970 range of heritability values on the observed scale, since one value is provided per neophobia 971 score level (i.e. five values in total); however, we could not derive repeatability estimates on 972 the observed scale for ordinal variables. 973

	Aggressiveness	Boldness	Aggressiveness
Correlation level	-	-	-
	Neophobia	Neophobia	Boldness
Phenotypic	-0.20*	-0.30*	0.02
	[-0.25; -0.12]	[-0.38; -0.23]	[-0.04; 0.10]
	0.00	0.00	-0.02
Between-individual	0.00	0.00	0.02

Table 2. Phenotypic, between- and within-individual, and additive genetic correlations 975 between aggressiveness, boldness and neophobia scores 976

978

Within-individual

Additive genetic

-0.28\*

0.00

[-0.38; -0.18]

[-0.06; 0.006]

Posterior modes and 95% credible intervals (CI) on the latent scale are shown. Asterisks 979 indicate estimates whose 95% CI do not encompass zero. 980

-0.43\*

-0.01

[-0.53; -0.32]

[-0.06; 0.06]

0.03

0.01

[-0.04; 0.12]

[-0.07; 0.06]

### 981 Table A1. Detailed description of the collared flycatcher pedigree from the Gotland

982 Island population

Pedigree statistics	
Records	2 218
Founders	1423
Maternities	728
Paternities	781
Mothers with $\geq 2$ offspring	120
Fathers with $\geq 2$ offspring	133
Full sibs	130
Maternal sibs	222
Maternal half sibs	92
Paternal sibs	230
Paternal half sibs	100
Maternal grandmothers	249
Maternal grandfathers	268
Paternal grandmothers	220
Paternal grandfathers	236
Maximum pedigree depth	15
Mean relatedness	5.29 x 10 <sup>-4</sup>

983

The pedigree statistics were obtained from all identified individuals involved in either aggressiveness, boldness or neophobia assays, and were extracted using the pedigreeStats and pedStatSummary functions from 'pedantics' R package (Morrissey & Wilson, 2010).

# **Table A2.** Output of the univariate models fitting aggressiveness score

	Number	r of aggressive behaviours/15	5 min					PC1					
	Withou	t fixed effect	With fix	ed effects				_ 101					
	Posterio	or 95% CI	Posterio mode	r 95% CI				Posterio	r 95% CI				
Fixed effects													
Intercept	1.22	[ 0.91 ; 1.50 ] *	-0.29	[ -0.72	; 0.	.26	]	-0.98	[ 1.30	;	- 0.65	]	*
Sex (male)			0.63	[ 0.44	; 0.	.81	] *	0.50	[ 0.35	;	0.62	]	*
Age (young)			0.23	[ -0.01	; 0.	.40	]	0.06	- [ 0.05	;	0.25	]	
Presence of tits (present	)		0.34	[ 0.19	; 0.	.50	] *	0.05	- [ 0.07	;	0.15	]	
Presence of flycatchers	(present)		0.27	[ -0.02	; 0.	.55	]	0.29	[ 0.05	;	0.49	]	*
Presence of the partner (	(present)		0.62	[ 0.44	; 0.	.80	] *	0.34	[ 0.20	;	0.44	]	*
Test number			-0.20	[ -0.27	; -0	).14	] *	-0.16	[ -	;	-	]	*

0.20 0.11

Dummy type (flycatcher)			0.08	[ -0.23	; 0.34	]	0.16	- [ ; 0.38 ] 0 0.01
Year (2012)			0.48	[ 0.08	; 0.87	] *	0.19	- [ ; 0.50 ] 0 0.05
Year (2013)			0.29	[ -0.09	; 0.67	]	0.10	- [ ; 0.46 ] 0 0.09
Sex*age (male*young)			0.28	[ 0.03	; 0.61	] *	0.24	[ 0.09 ; 0.50 ] * 0
Random effects								
$V_{ m A}$	0.00	[ 0.00 ; 0.32 ]	0.00	[ 0.00	; 0.41	]	0.00	[ 0.00 ; 0.16 ] 0
$V_{ m PE}$	0.60	[ 0.28 ; 0.80 ] *	0.52	[ 0.11	; 0.71	] *	0.19	[ 0.03 ; 0.28 ] * 1
$V_{\mathbf{M}}$	0.00	[ 0.00 ; 0.20 ]	0.00	[ 0.00	; 0.17	]	0.00	[ 0.00 ; 0.12 ] 0
$V_{ m F}$	0.00	[ 0.00 ; 0.20 ]	0.00	[ 0.00	; 0.24	]	0.23	[ 0.06 ; 0.37 ] * 0
$V_{ m plot}$	0.07	[ 0.02 ; 0.16 ] *	0.05	[ 0.01	; 0.15	] *	0.01	[ 0.00 ; 0.04 ] 0
$V_{ m decoy}$	0.00	[ 0.00 ; 0.03 ]	0.00	[ 0.00	; 0.04	]	0.00	[ 0.00 ; 0.03 ] 0
$V_{ m song}$	0.02	[ 0.00 ; 0.09 ]	0.01	[ 0.00	; 0.09	]	0.00	[ 0.00 ; 0.03 ] 0

$V_{\sf obs}$	0.25	[ 0.14 ; 0.55 ] *	0.30	[ 0.15 ; 0.63 ] *	0.06 [ 0.02 ; 0.16 ] * 0
Vε	2.56	[ 0.91 ; 1.50 ] *	2.11	[ 1.97 ; 2.34 ] *	1.49 [ 1.40 ; 1.57 ] * 1
Derived estimates					
R <sub>latent</sub>	0.18	[ 0.15 ; 0.23 ] *	0.18	[ 0.13 ; 0.22 ] *	0.23 [ 0.17 ; 0.27 ] * 0
$R_{ m obs}$	0.03	[ 0.02 ; 0.04 ] *	0.03	[ 0.02 ; 0.04 ] *	0.21 [ 0.16 ; 0.25 ] *
$h^2_{\text{latent}}$	0.00	[ 0.00 ; 0.08 ]	0.00	[ 0.00 ; 0.12 ]	0.00 [ 0.00 ; 0.08 ] 0
$h^2_{ m obs}$	0.00	[ 0.00 ; 0.01 ]	0.00	[ 0.00 ; 0.01 ]	0.00 [ 0.00 ; 0.07 ] P
N		4680		3271	3271
Effective sample size		> 2357		> 2234	> 2209

Models for the general aggressiveness score, based on the number of aggressive behaviours standardized per 15 min, are shown without and with fixed effects. We also present the posterior modes and 95% credible intervals (CI) of models for alternative aggressiveness scores: the first axis of the principal component analysis (PC1) and the discrete score (see Appendix). Asterisks indicate estimates whose 95% CI do not encompass zero. For categorical fixed terms, estimates refer to the category indicated in parentheses.  $V_A$ ,  $V_{PE}$ ,  $V_M$ ,  $V_F$  and  $V_e$  refer to the additive genetic, permanent environment, maternal, paternal and residual variances, respectively.  $V_{plot}$ ,  $V_{decoy}$ ,  $V_{song}$ ,  $V_{obs}$  refer to the variances associated with the plot, the decoy set used, the song track played and observer identity, respectively. N is the sample size of the data set used in the model.

- 9 Repeatability and heritability estimates are given both on the latent scale ( $R_{latent}$ ,  $h^2_{latent}$ ) and on the observed scale ( $R_{obs}$ ,  $h^2_{obs}$ ) whenever these
- 10 could be estimated.

# 11 Table A3 Output of the models fitting boldness scores.

	Maxim	um latency-indiv	idual latency	I		Latency/feeding rate	Discrete sc
	Withou	t fixed effects		Without	it fixed effects		Discrete se
	Posterio	or 95% CI		Posterio	or 95% CI	Posterior 95% CI mode	Posterior mode
Fixed effects							
Intercept	8.00	[ 7.97 ; 8	3.04 ] *	8.05	[ 7.96 ; 8.10 ] *	-3.78 [ -4.07 ; ] * 3.46	4.80
Sex (male)				-0.04	[ -0.12 ; 0.02 ]	-0.10 [ -0.29 ; 0.13 ]	-1.33
Age (young)				-0.02	[ -0.10 ; 0.05 ]	-0.12 [ -0.39 ; 0.13 ]	-0.65
Year (2012)				-0.03	[ -0.10 ; 0.04 ]	0.05 [ -0.19 ; 0.43 ]	-0.38
Year (2013)				0.08	[ 0.00 ; 0.16 ] *	0.47 [ 0.10 ; 0.73 ] *	1.46
No. of chicks				0.03	[ 0.00 ; 0.06 ] *	0.26 [ 0.15 ; 0.33 ] *	0.36
Sex*age (male*you	ng)			0.03	[ -0.08 ; 0.13 ]	0.30 [ -0.18 ; 0.55 ]	0.60
Random effects							
$V_{\mathbf{A}}$	0.00	[ 0.00 ; 0	).02 ]	0.00	[ 0.00 ; 0.03 ]	0.00 [ 0.00 ; 0.20 ]	0.01

$V_{\rm PE}$	0.00	[	0.00	;	0.02	]	0.00	[	0.00	;	0.03	]	0.00	[	0.00	;	0.25	]		0.04
$V_{\mathrm{M}}$	0.00	[	0.00	;	0.02	]	0.00	[	0.00	;	0.02	]	0.00	[	0.00	;	0.13	]		0.01
$V_{ m F}$	0.00	[	0.00	;	0.01	]	0.00	[	0.00	;	0.01	]	0.00	[	0.00	;	0.15	]		0.01
$V_{ m plot}$	0.00	[	0.00	;	0.00	]	0.00	[	0.00	;	0.01	]	0.00	[	0.00	;	0.07	]		0.00
$V_{\rm obs}$	0.00	[	0.00	;	0.00	]	0.00	[	0.00	;	0.00	]	0.00	[	0.00	;	0.03	]		0.00
νε	0.16	[	0.14	;	0.17	] *	0.13	[	0.11	;	0.16	] *	1.20	[	0.93	;	1.37	]	*	10.00
Derived estimates																				
R <sub>latent</sub>	0.11	[	0.01	;	0.21	] *	0.19	[	0.07	;	0.31	] *	0.12	[	0.01	;	0.29	]	*	0.28
$R_{ m obs}$	0.10	[	0.01	;	0.19	] *	0.12	[	0.02	;	0.23	] *	0.04	[	0.01	;	0.14	]	*	
$h^{2}_{latent}$	0.00	[	0.00	;	0.10	]	0.00	[	0.00	;	0.18	]	0.00	[	0.00	;	0.14	]		0.00
$h^2_{ m obs}$	0.00	[	0.00	;	0.09	]	0.00	[	0.00	;	0.09	]	0.00	[	0.00	;	0.06	]		Posterior m
Ν		10	)64					9	014					64	11					
Effective sample size															>					
Effective sample size		>	1718					>	× 1979						2053					

12 Models for the general boldness estimate (maximum latency to enter the nestbox after human disturbance observed in the entire data set minus

13 the individual latency) are shown without and with fixed effects. We also present the posterior modes and 95% credible intervals (CI) of models

for alternative boldness estimates: the log-transformed and inverse ratio of the latency to enter the nestbox after human disturbance over the 14 feeding rate during the time remaining and the discrete score based on the latency (see Appendix). Asterisks indicate estimates whose 95% CI do 15 not encompass zero. For categorical fixed terms, estimates refer to the category indicated in parentheses. VA, VPE, VM, VF and VE refer to the 16 additive genetic, permanent environment, maternal, paternal and residual variances, respectively. V<sub>plot</sub> and V<sub>obs</sub> refer to the variances associated 17 18 with the plot and the observer identity, respectively. N is the sample size of the data set used in the model. Repeatability and heritability estimates are given both on the latent scale ( $R_{latent}$ ,  $h^2_{latent}$ ) and on the observed scale ( $R_{obs}$ ,  $h^2_{obs}$ ) whenever these could be estimated. Repeatability and 19 heritability estimates for the latency/feeding rate model (fitted with a Gaussian distribution) differ between the latent and observed scales because 20 we accounted for the inverse log transformation in the QGicc and QGparams functions ('QGglmm' R pachage; de Villemereuil et al. 2016). 21

# 22 Table A4. Output of the models fitting neophobia scores

	Discrete so	core							Binary							I
	Without fi	ixed effect		With fixed	d effects				_ Dillal y							J
	Posterior mode	95% CI		Posterior mode	95% CI				Posterior mode		6 CI					]
Fixed effects																
Intercept	4.22	[ 3.40	; 4.94 ] *	5.08	[ 3.63	;	6.38	] *	-1.67	[ -	3.39	;	-0.30	]	*	(
Sex (male)				-1.89	[ -2.78	;	-1.07	] *	-1.45	[ -	2.44	;	-0.41	]	*	-
Age (young)				-0.18	[ -1.01	;	0.76	]	0.07	[ -	1.26	;	0.92	]		
Year (2012)				1.62	[ 0.72	;	2.93	] *	2.19	[ (	).59	;	3.85	]	*	
Year (2013)				1.21	[ -0.05	;	2.37	]	1.84	[-	0.06	;	3.34	]		
Boldness				-1.58	[ -1.93	;	-1.12	] *	-1.34	[ -	1.95	;	-0.93	]	*	

No. of chicks								-0.71	[	-1.03	;	-0.37	]	*	-0.71	[	-1.15	;	-0.30	]	*	-0
Sex*age (male*young)								-1.02	[	-2.44	;	0.11	]		-0.79	[	-2.75	;	0.69	]		-0
Random effects																						
$V_{ m A}$	0.02	[	0.00	;	2.90	]		0.03	[	0.00	;	5.76	]		0.08	[	0.00	;	8.41	]		0.
$V_{PE}$	0.04	[	0.00	;	5.77	]		0.06	[	0.00	;	9.03	]		0.05	[	0.00	;	10.70	]		0.
$V_{\mathbf{M}}$	0.04	[	0.00	;	6.49	]		0.06	[	0.00	;	8.81	]		0.12	[	0.00	;	22.53	]		0.
$V_{ m F}$	0.02	[	0.00	;	5.75	]		0.03	[	0.00	;	7.85	]		0.04	[	0.00	;	11.33	]		0.
$V_{plot}$	0.98	[	0.26	;	2.50	]	*	1.01	[	0.27	;	3.42	]	*	0.80	[	0.00	;	3.08	]		0.
$V_{\sf obs}$	0.28	[	0.02	;	1.01	]	*	0.38	[	0.00	;	1.41	]		0.80	[	0.14	;	3.14	]	*	0.
Vε	10.00	[	10.00	;	10.00	]		10.00	[	10.00	;	10.00	]		10.00	[	10.00	;	10.00	]		0.
Derived estimates																						
R <sub>latent</sub>	0.39	[	0.25	;	0.54	]	*	0.50	[	0.33	;	0.63	]	*	0.56	[	0.35	;	0.75	]	*	0.
$R_{ m obs}$															0.29	[	0.17	;	0.48	]	*	0.
$h^2_{\text{latent}}$	0.00	[	0.00	;	0.15	]		0.00	[	0.00	;	0.22	]		0.00	[	0.00	;	0.26	]		0.

$h^2_{ m obs}$	Posterior mode range: 0.00-0.00	Posterior mode range: 0.00-0.00	0.00	[ 0.00 ; 0.14 ]	C
Ν	1 064	914		914	
Effective sample size	> 2302	> 2317		> 1583	

0.

Models for the general neophobia discrete score, based on the latency to enter the nestbox in presence of the novel object, are shown without and 23 with fixed effects We also present the posterior modes and 95% credible intervals (CI) of models for alternative neophobia estimates: a binary 24 variable (individual entered versus did not enter in the presence of the novel object) and a continuous latency to enter the nestbox, for individuals 25 that entered in the presence of the novel object (see Appendix). Asterisks indicate estimates whose 95% CI do not encompass zero. For 26 categorical fixed terms, estimates refer to the category indicated in parentheses. VA, VPE, VM and VF refer to the additive genetic, permanent 27 environment, maternal and paternal variances, respectively. The residual variance  $V_{\varepsilon}$  was set to 10.  $V_{\text{plot}}$  and  $V_{\text{obs}}$  refer to the variances associated 28 with the plot and the observer identity, respectively. N is the sample size of the data set used in the model. Repeatability and heritability estimates 29 are given both on the latent scale ( $R_{latent}$ ,  $h^2_{latent}$ ) and the observed scale ( $R_{obs}$ ,  $h^2_{obs}$ ) whenever these could be estimated. 30

**Table A5.** Output of the trivariate model fitting aggressiveness, boldness and neophobia scores with an individual random effect

	Posterior	94	5% CI				
	mode						
Fixed effects							
Aggressiveness	1.73	[	1.65	;	1.82	]	*
Boldness	7.97	[	7.93	;	8.02	]	*
Neophobia	4.44	[	3.79	;	5.00	]	*
Sex (male)	-0.02	[	-0.06	;	0.04	]	
Year (2012)	0.04	[	-0.02	;	0.09	]	
Year (2013)	0.09	[	0.03	;	0.15	]	*
Random effects							
$V_{ m ind, \ aggressiveness}$	0.42	[	0.16	;	0.66	]	*
$V_{\mathrm{ind, \ boldness}}$	0.02	[	0.00	;	0.04	]	
$V_{ m ind,\ neophobia}$	6.14	[	2.96	;	10.27	]	*
$V_{arepsilon}$ , aggressiveness	1.36	[	1.09	;	1.62	]	*

$V_{arepsilon,  ext{ boldness}}$	0.13	[	0.12	;	0.16	]	*
$V_{arepsilon,  { m neophobia}}$	10.00	[	10.00	;	10.00	]	
Covind, aggressiveness-neophobia	0.00	[	-0.09	;	0.11	]	
Covind, aggressiveness-boldness	0.00	[	-0.01	;	0.01	]	
Covind, boldness-neophobia	0.00	[	-0.03	;	0.02	]	
$Cov_{\epsilon,aggressiveness-neophobia}$	-1.03	[	-1.37	;	-0.63	]	*
$Cov_{\epsilon, aggressiveness-boldness}$	0.02	[	-0.02	;	0.05	]	
$Cov_{\epsilon, \text{ boldness-neophobia}}$	-0.49	[	-0.62	;	-0.36	]	*
Ν		16	589				
Effective sample size:		>	1740				

35 V stands for variance terms and Cov for covariance terms (posterior mode and 95% credible interval (CI)). The individual effect was the only

36 random term included; 'ind' and 'ɛ' stand for between-individual and residual terms, respectively. The residual variance for neophobia score was

37 fixed to 10 (see text for the distributions used for the three scores). Asterisks indicate estimates whose 95% CI do not encompass zero.

**Table A6.** Output of the trivariate model fitting aggressiveness, boldness and neophobia scores with an additive genetic effect

	Posterior	94	5% CI				
	mode	95% CI					
<b>Tixed effects</b>							
Aggressiveness	1.71	[	1.62	;	1.79	]	*
Boldness	7.97	[	7.93	;	8.02	]	*
Neophobia	4.12	[	3.55	;	4.62	]	*
Sex (male)	-0.01	[	-0.05	;	0.04	]	
Year (2012)	0.04	[	-0.02	;	0.09	]	
Year(2013)	0.09	[	0.03	;	0.15	]	*
andom effects							
$V_{ m A, aggressiveness}$	0.15	[	0.00	;	0.38	]	
$V_{ m A, \ boldness}$	0.02	[	0.00	;	0.04	]	
$V_{ m A,\ neophobia}$	4.81	[	2.07	;	7.52	]	*
$V_{arepsilon,  ext{ aggressiveness}}$	1.46	[	1.29	;	1.78	]	*

	$V_{\epsilon, \text{ boldness}}$	0.14	[	0.12	;	0.16	]	*
	$V_{arepsilon,  { m neophobia}}$	10.00	[	10.00	;	10.00	]	
	$\mathrm{Cov}_{\mathrm{A},\ \mathrm{aggressiveness-neophobia}}$	0.00	[	-0.06	;	0.07	]	
	$Cov_{A, aggressiveness-boldness}$	0.00	[	0.00	;	0.00	]	
	$Cov_{A, boldness-neophobia}$	0.00	[	-0.02	;	0.02	]	
	$Cov_{\epsilon, aggressiveness-neophobia}$	-1.01	[	-1.35	;	-0.64	]	*
	$Cov_{\epsilon,aggressiveness\text{-boldness}}$	0.01	[	-0.02	;	0.05	]	
	$Cov_{\epsilon, \text{ boldness-neophobia}}$	-0.48	[	-0.61	;	-0.37	]	*
N			16	686				
E	ffective sample size		>	1816				

V stands for variance terms and Cov for covariance terms (posterior mode and 95% credible interval (CI)). The additive genetic effect was the only random term included; 'A' and ' $\varepsilon$ ' stand for additive genetic and residual terms, respectively. The residual variance for neophobia score was fixed to 10 (see text for the distributions used for the three scores). Asterisks indicate estimates whose 95% CI do not encompass zero. Positive additive genetic variances are found here because permanent environment and parental effects are not taken into account.

- 1 Table A7. Output of a principal component analysis of behaviours recorded during the
- 2 aggressiveness assays

	Coordinates		Contribution		
	PC1	PC2	PC1	PC2	
Moves < 2 m from the nestbox	0.83	-0.12	36.12	0.95	
Stationary flights towards the decoy	0.73	-0.21	27.74	2.99	
Attacks towards the decoy	0.67	-0.24	23.24	4.06	
Moves between 2 and 5 m from the nestbox	0.42	0.65	9.40	29.96	
Moves between 5 and 10 m from the nestbox	-0.03	0.74	0.03	38.49	
Chases of live birds	0.26	0.58	3.47	23.55	

4 We used the function PCA from the 'FactoMineR' R package (Lê, Josse, & Husson, 2008).

### 5 FIGURE CAPTIONS

6

Figure A1. Sex differences in aggressiveness, boldness and neophobia scores (means ± 95% confidence interval). (a) Between-sexes differences in average aggressiveness score for a given individual in a given year depending on age (yearling versus older). (b) Between-sexes differences in the latency to return after human disturbance for the period without a novel object, as a proxy of (inverse) boldness, and for the period with the novel object, as a proxy of neophobia. Number of observations is indicated near each estimate.

13

Figure A2. (a) Aggressiveness and (b) boldness scores depending on the neophobia score (means ± 95% confidence interval). See text for the definitions of the scores. Aggressiveness is here the averaged value of all scores for a given individual in a given year. Number of observations is indicated near each estimate.

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Figure A3. Distribution of the discrete scores for (a) aggressiveness, (b) boldness and (c)neophobia. Number of observations is indicated above each bar.









