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ABSTRACT

Personality traits and their correlations have been shown to be linked with life history strategies and fitness in various species. Between-individual correlations (i.e. behavioural syndromes) between personality traits can affect the evolutionary responses of these traits to environmental variation. Understanding the genetic and ecological determinants of personality traits and their interactions as behavioural syndromes in the wild is thus needed to shed light on the mechanisms shaping their evolution. Partitioning the observed (co)variance in these traits, however, requires large numbers of repeated behavioural measures on many individuals of known relatedness level. In the absence of such data, it is thus often assumed that phenotypic (co)variances inform about (1) underlying between-individual (co)variances (i.e. ignoring within-individual (co)variances) and (2) underlying genetic (co)variances. We tested these assumptions using three personality traits collected during 3 years on a long-term monitored breeding population of collared flycatchers, *Ficedula albicollis*. We partitioned the observed phenotypic (co)variance of aggressiveness, boldness and neophobia into genetic, permanent environment and parental components, and we estimated the repeatability and heritability of these traits and their between-individual correlations. All three traits were repeatable between years (at least on the latent scale) but none were heritable. Permanent environment effects explained 15% of the phenotypic variance in aggressiveness, and parental effects explained 25% of the phenotypic variance in neophobia, in line with previous studies in wild populations. The three traits showed phenotypic correlations but no between-individual correlations and no additive genetic covariance. Thus, our results did not support the assumptions that phenotypic covariance reflects behavioural syndromes and genetic covariance. We discuss the reasons for the absence of heritability and between-individual and

genetic covariance between these three personality traits in light of the possible selective pressures acting on this population.

INTRODUCTION

Over the past two decades, personality traits, that is, repeatable between-individual behavioural differences across time and contexts (Réale, Reader, Sol, McDougall, & Dingemanse, 2007), have received increasing attention in animal behavioural and evolutionary ecology studies (Bell, 2007; Bell, Hankison, & Laskowski, 2009; Carere & Maestripieri, 2013; Dingemanse & Réale, 2005; Réale et al., 2007; Sih, Bell, & Johnson, 2004; Sih, Bell, Johnson, & Ziemba, 2004). Five ecologically important personality axes have been identified to characterize the behavioural responses of individuals when interacting with their environment (activity, exploration, boldness) and with others (aggressiveness, sociability; Réale et al., 2007). Personality traits have been shown to depend on ecological parameters (e.g. Réale et al., 2007; Sih, Cote, Evans, Fogarty, & Pruitt, 2012), to be heritable (e.g. van Oers, de Jong, van Noordwijk, & Drent, 2005; van Oers & Sinn, 2013), to be linked to life history traits or fitness (Dingemanse, Both, Drent, & Tinbergen, 2004; Dingemanse & Réale, 2013; Duckworth & Kruuk, 2009; Reale et al., 2010; Schuett, Tregenza, & Dall, 2010; Smith & Blumstein, 2008) and often to correlate with each other at the individual level, forming so-called behavioural syndromes (Garamszegi, Markó, & Herczeg, 2012; Andrew Sih, Bell, & Johnson, 2004; van Oers & Sinn, 2013). Such correlations may result from a functional integration of personality traits favoured by selection when interactions between these traits increase individuals' fitness in given environmental conditions (e.g. Dingemanse et al., 2007). Altogether, these various results reveal the crucial role that personality traits may

play in shaping evolutionary processes in wild populations (Dingemanse et al., 2004; Duckworth & Badyaev, 2007; Karlsson Green, Eroukhmanoff, Harris, Pettersson, & Svensson, 2016; Niemelä, Lattenkamp, & Dingemanse, 2015).

Understanding the evolution of personality traits and their associations in behavioural syndromes requires understanding the mechanisms underlying these between-individual differences in behaviour and their interactions, including their genetic basis. Phenotypic correlations between personality traits result from the addition of between-individual correlations (defining behavioural syndromes per se, whether genetic or nongenetic) and within-individual (or residual) correlations (Brommer, 2013; Dingemanse & Réale, 2013; Dingemanse, Kazem, Réale, & Wright, 2010). Assessing the relative contribution of between- and within-individual correlation components in observed phenotypic correlations can be crucial because between-individual correlations may constrain the independent evolution of the traits involved and thus may have major evolutionary consequences in the wild (Sih, Bell, & Johnson, 2004).

To reliably quantify this relative contribution of between- and within-individual correlations, multiple measurements of the personality traits considered must be collected on a large number of individuals. When only single measurements are available, it is often assumed that a phenotypic correlation observed between personality traits reflects an underlying between-individual correlation, i.e. the within-individual correlation is negligible (the so-called 'individual gambit', Brommer, 2013). Furthermore, to quantify genetic variance in personality traits and genetic covariance between them, the level of genetic relatedness between individuals measured has to be incorporated (e.g. via pedigree information) in the models. When relatedness information is unavailable, it is often assumed that the observed phenotypic (co)variance reflects the underlying genetic (co)variance (the so-called

99 'phenotypic gambit'; Grafen, 1984; Hadfield, Nutall, Osorio, & Owens, 2007; van Oers &
100 Sinn, 2011).

101 These two crucial assumptions have been tested empirically in various species, and
102 recent meta-analyses including over 30 studies, among which 25 are from wild populations,
103 have confirmed their overall validity (Brommer & Class, 2017; Dochtermann, 2011;
104 Dochtermann, Schwab, & Sih, 2015). Across these studies, the sign (and to a certain extent
105 the magnitude) of the phenotypic correlations between personality traits reliably informed on
106 the sign (and the magnitude) of the between-individual correlations (Brommer & Class, 2017)
107 and of the genetic correlations (Dochtermann, 2011). Furthermore, 52% of the between-
108 individual variation in personality traits taken separately was explained by additive genetic
109 variance (Dochtermann et al., 2015). A recent empirical study on a wild population of yellow-
110 bellied marmots, *Marmota flaviventris*, estimated the proportion of phenotypic (co)variance
111 explained by genetic, permanent environment and maternal (co)variances between four
112 different personality traits: docility, exploration, activity and sociability (Petelle, Martin, &
113 Blumstein, 2015). Results showed additive genetic variations, as well as maternal and
114 permanent environment variations, in all four traits and a positive genetic correlation between
115 activity and sociability (Petelle et al., 2015). More of such integrative studies partitioning the
116 observed phenotypic (co)variance in multiple personality traits simultaneously are needed in
117 different biological models with contrasting life histories to better understand the mechanisms
118 underlying and possibly constraining the evolution of correlated personality traits.

119 In this study, we assessed the genetic basis of three personality traits, together with the
120 relative contribution of between- and within-individual variations in, and correlations
121 between, these traits, chosen for their potentially important effects on crucial ecological
122 processes (here, nest site acquisition and defence against competitors and predators) in a
123 natural population of a small territorial, short-lived, migrant passerine bird, the collared

124 flycatcher, *Ficedula albicollis*. During 3 consecutive years, we measured for several hundreds
125 of breeding pairs in the field (1) aggressiveness towards competitors (as the agonistic reaction
126 to simulated territorial intrusions by intra- and interspecific competitors), (2) boldness
127 towards predators (as the latency to resume nestling feeding after human disturbance) and (3)
128 neophobia, possibly reflecting exploration (as the latency to resume nestling feeding in the
129 presence of a novel object on the nest site; following Réale et al.'s 2007 definitions). To
130 identify the mechanisms underlying the phenotypic (co)variation in these behavioural traits,
131 we assessed to what extent (1) additive genetic, parental or permanent environment effects
132 contributed to the observed phenotypic (co)variance, while accounting for fixed individual
133 (sex, age) covariates, and (2) between-individual correlations explained phenotypic
134 correlations between these three traits. Based on many previous studies on personality traits in
135 populations of passerines of similar ecology (e.g. Dingemanse, Both, Drent, van Oers, & van
136 Noordwijk, 2002; Drent, Oers, & Noordwijk, 2003; Duckworth & Badyaev, 2007;
137 Garamszegi et al., 2015; Garamszegi, Rosivall, et al., 2012; van Oers, Drent, de Goede, & van
138 Noordwijk, 2004), we expected heritable differences to partly explain variation in
139 aggressiveness, boldness and neophobia in our study population. Furthermore, high
140 aggressiveness, high boldness and low neophobia may allow individuals to efficiently secure
141 and defend a breeding site when they are unfamiliar with the environment (e.g. for dispersers:
142 Cote, Clobert, Brodin, Fogarty, & Sih, 2010; Duckworth & Kruuk, 2009; Korsten, van
143 Overveld, Adriaensen, & Matthysen, 2013). Thus we predicted a functional integration and
144 (possibly genetically based) between-individual correlations between these traits. In another
145 population of collared flycatchers, male aggressiveness and boldness, but not neophobia, were
146 found to be phenotypically correlated in some years (Garamszegi et al., 2015; Garamszegi,
147 Eens, & Török, 2009). This population and ours, however, differ greatly in both demographic
148 functioning (e.g. male age structure: Hegyi, Rosivall, & Török, 2006) and selective pressures

149 (e.g. nest predation: Doligez & Clobert, 2003; sexual selection: Qvarnström, 1997; see also
150 Rosivall, Török, Hasselquist, & Bensch, 2004), which may affect the fitness consequences of
151 interactions between personality traits, and thus their potential functional integration.
152
153
154 **METHODS**
155
156 ***Study species and population monitoring***
157 Collared flycatchers are migratory cavity nesters and readily breed in artificial nestboxes,
158 providing easy access to parents' identity and breeding data. Between 2011 and 2013, we
159 conducted the behavioural tests (see below) on 1131 pairs breeding in nestboxes spread over
160 14 to 22 forest patches in our study population located on the island of Gotland (Sweden,
161 Baltic Sea). Each year since 1980, nests in boxes have been monitored at least weekly from
162 late April until early July, allowing us to record major breeding variables (laying and hatching
163 dates; clutch size; nestling number, condition and fledging success). Breeding pairs were
164 captured, identified and ringed if previously unringed; females were caught during incubation
165 and males while feeding nestlings. Nestlings were ringed between day 8 and day 13 after
166 hatching; fledging typically occurs 16 days after hatching. Adult and nestling identification
167 every year combined with a high return rate of both adults (approximately 40%) and juveniles
168 (approximately 10%) for such a short-lived passerine bird (Gustafsson, 1989) allowed us to
169 establish a high-quality social pedigree of the population which has previously been used in
170 several quantitative genetic studies (e.g. Evans & Gustafsson, 2017; Merilä & Sheldon, 2000;
171 Sheldon, Kruuk, & Merilä, 2003; Appendix Table A1). In this population, approximately 15%
172 of all nestlings are extrapair (Sheldon & Ellegren, 1999), a percentage considered low enough
173 for quantitative genetic models to provide valid (i.e. only slightly underestimated) heritability

estimates from the social pedigree (Charmantier & Réale, 2005; Firth, Hadfield, Santure, Slate, & Sheldon, 2015), even though no information is available yet on how extrapair paternity may affect genetic covariances. The clear sexual dimorphism in plumage coloration in this species allowed an easy discrimination of adult males (black and white plumage with a white forehead patch) from females (brown plumage; Svensson, 1992), even from several metres away during behavioural tests. Adults could also be aged by plumage criteria (yearlings versus older adults; Svensson, 1992).

Aggressiveness score

We measured the level of aggressiveness of breeding flycatchers soon after settlement, during nest building or early laying, that is, when the risk of losing a nestbox to a competitor is highest in this single-clutch species. During the breeding season, collared flycatchers compete for nest sites with conspecifics but also with great tits, *Parus major*, the second most abundant species breeding in nestboxes in the study area (Gustafsson, 1987). Aggressiveness towards conspecific intruders was shown to decrease after the start of incubation (Král & Bícík, 1989) even though aggressiveness towards great tit intruders remained high throughout the breeding cycle (Král & Bícík, 1992). To elicit an aggressive response from a focal flycatcher pair, we simulated the intrusion of competitors at the nest of the pair by attaching to its nestbox clay decoys mimicking either a flycatcher pair or a single (male) great tit. We used a pair (one male and one female) for flycatcher decoys to elicit and measure an aggressive response by both pair members, that is, to avoid a sex-specific response towards this intraspecific stimulus, while the response to the interspecific stimulus (great tit decoy) was not expected to differ depending on the sex of the decoy. In addition, we simultaneously broadcast male songs corresponding to the decoy(s) species with a loudspeaker placed just under the nestbox. To avoid pseudoreplication, we randomly used one of eight different sets of decoys and one of

five different song tracks per species for each test. After attaching the decoys to the nestbox and the loudspeaker under the box, the observer sat under a camouflage net approximately 8–10 m away from the nestbox and recorded the following behaviours for each pair member: (1) movements between perches and perching position (within 2 m, between 2 and 5 m or between 5 and 10 m away from the nestbox), (2) agonistic behaviours towards a decoy (attacks and stationary flights in front of the decoy) and (3) chases towards living birds attracted by the stimulus. A behavioural test started with an observation period of 15 min but we lengthened the test by up to 5 min when an individual arrived between 10 and 15 min after the start of the test, and up to 5 additional min if its partner arrived during this extra time, so that we could observe the behavioural response of each pair member for at least 5 min. Each test thus lasted between 15 and 25 min. If an individual was observed during less than 5 min before the end of the test, it was discarded from the analyses and these observations were therefore not used later on.

Aggressiveness level was measured as the sum of the number of movements within 2 m of the nestbox, attacks, stationary flights and chases. We included this latter behaviour because chasing a live intruder may have prevented the focal flycatcher from interacting with the decoy, while reflecting an aggressive territory defence response. The number of each type of behavioural response (movements, agonistic behaviours and chases) was standardized by the time interval between the first observation of the individual and the end of the test, rescaled to 15 min. Using alternative scores did not qualitatively change the results (see Appendix and Table A2). We conducted aggressiveness tests two to four times per focal pair over a 5-day interval, with at most one test per day and tests on 2 days in a row. The stimulus used (intra- / interspecific decoys) was alternated between tests after a random assignment for the first test. An aggressiveness score was computed for each individual for each test. We obtained aggressiveness responses (for more than 5 min at least once per year) for 1974

224 individuals (including unidentified ones; 961 females and 1014 males in 1046 nests). Among
225 those, 825 females and 667 males were later captured and identified, and thus used for
226 heritability estimation. In 601 breeding pairs both partners were identified and responded to
227 the tests and in 273 only one partner responded. We obtained repeated estimates for 502 and
228 445 identified females and males, respectively.
229
230 ***Boldness and neophobia scores***
231 During nestling rearing, we estimated (1) boldness level by measuring the individual's
232 reaction towards the presence of a human observer near the nestbox and (2) neophobia level
233 by measuring the reaction towards the presence of a novel object on the nestbox (i.e. in a
234 familiar environment), following the definitions from Réale et al. (2007). As advised in
235 Greenberg and Mettke-Hofmann (2001), we measured our behavioural responses as the
236 latency to perform a highly motivating action (here, feeding their nestlings) after disturbance.
237 We conducted one combined boldness/neophobia test per breeding pair when the nestlings
238 were 5 days old, that is, at the beginning of the period of highest provisioning by parents (and
239 before we caught the parents to avoid any behavioural interference). A test consisted of two
240 consecutive periods of approximately 1 h each: the behaviour of the parents was recorded first
241 without any change in the surroundings of the nestbox, that is, without the novel object, and
242 second with a novel object (here a coloured figurine approximately 7 cm high) attached near
243 the entrance hole of the nestbox. Both periods were video-recorded from a distance (6–8 m).
244 At the beginning of each period, the observer checked the camouflaged video recorder,
245 walked to the nestbox, opened it to check nestling satiety, closed it, and then left the area. The
246 test was abandoned if the nestlings were very hungry to avoid them starving if the parents
247 were too disturbed by the test.

We estimated boldness score using the latency to enter the nestbox after the departure of the observer in the first period (i.e. without the novel object). Reaction to disturbance by humans has previously been used in boldness tests in this species (e.g. Garamszegi et al., 2009). To ease interpretation (i.e. increasing values of boldness score corresponding to increasing level of boldness), we transformed the latency to enter the nestbox such that the boldness score of an individual was the maximum latency observed in the entire data set minus the latency for this individual. We estimated neophobia score based on the latency to enter the box after the departure of the observer in the second period (i.e. in the presence of the novel object). For both boldness and neophobia scores, individuals that did not enter the nestbox during the first period of the test were not used in the analysis (187 of 1251 observations, i.e. 15%). Individuals that entered the nestbox during the first but not the second part of the test (411 of 1064 observations, i.e. 39%) were considered as the most neophobic ones but could not be assigned a latency. To include them in the analyses, we discretized the latency to enter the nestbox in the second period into four categories based on its quartiles, assigning values from 1 to 4 for increasing latencies, and adding a fifth category including individuals that did not enter in the second part of the test. Using alternative scores for boldness and neophobia did not qualitatively change the results (see Appendix and Tables A2 and A3). We obtained boldness and neophobia estimates for 849 identified individuals (472 females and 378 males). Over the 3 years, we assessed 403 unique breeding pairs where both identified partners responded to the tests and 185 pairs where only one partner responded. We obtained repeated boldness and neophobia scores (i.e. several years in a row) for 66 females and 65 males.

Repeatability and heritability of aggressiveness, boldness and neophobia

We estimated the repeatability of aggressiveness, boldness and neophobia scores as well as their heritability by fitting three separate univariate animal models. The models included the following random effects: additive genetic effect (associated with the pedigree), individual identity for the repeated measures per individual (permanent environment effect), maternal and paternal identities, forest patch, observer identity (the person observing and reporting the behaviours onsite for the aggressiveness tests and the person extracting latencies from the video recording for the boldness and neophobia tests). The models of the aggressiveness score also included the broadcast song track and decoy set identifiers. In addition, the models included the following fixed effects, to control for potential confounding factors: sex, age (two levels: yearling versus older) and their interaction, as well as year (three levels: 2011, 2012, 2013). The aggressiveness model also included stimulus type (two levels: flycatcher versus great tit decoys), the order of the test within a year (continuous variable: first to fourth), the presence of the partner during the test (binary variable: yes/no) and the presence of other live flycatchers or great tits (binary variable: yes/no). The boldness and neophobia models included the number of ringed nestlings as a proxy of the motivation to enter the nestbox to feed nestlings. The neophobia models included the boldness score to control for the effect of the human disturbance at the beginning of the period with the novel object. All continuous fixed terms were centred and standardized prior to analysis to allow comparisons between effects.

Repeatabilities (R) were estimated as the ratio of the sum of the additive genetic (V_A), permanent environment (V_{PE}), maternal and paternal identities variances (V_M and V_F, respectively) over the total phenotypic variance (V_P, sum of all variances; Falconer & Mackay, 1996; with possibly an additional term accounting for the distribution variance, Nakagawa & Schielzeth, 2010). Narrow-sense heritabilities (h^2) were estimated as the ratio of the additive variance V_A over the phenotypic variance V_P. The presence of fixed effects in

297 models did not result in over- or underestimating repeatability and heritability estimates (as
298 warned against in Nakagawa & Schielzeth, 2010), because estimates were similar when
299 obtained from models with only the intercept (presented in the main text) and from models
300 with the previously described fixed effects (see Appendix Tables A1, A2, A3; except for a
301 slight overestimation of the boldness and neophobia repeatabilities on the latent scale). For
302 aggressiveness, we also estimated within-year repeatability by replacing the additive genetic
303 and permanent environment effects by a unique identifier per individual per breeding season.
304
305 ***Correlations between aggressiveness, boldness and neophobia scores***
306 We estimated the between- and within-individual correlations between aggressiveness,
307 boldness and neophobia scores by fitting a trivariate mixed-effects model (Dingemanse &
308 Dochtermann, 2013). For this model, aggressiveness was averaged over all aggressiveness
309 scores of an individual in a given year (i.e. over up to four estimates). We chose this
310 averaging approach because aggressiveness score (1) differed depending on decoy species,
311 chosen at random for the first test, and (2) decreased with the order of the test (i.e. due to
312 habituation). Averaging all aggressiveness scores of an individual in a given year allowed us
313 to control for these differences and decrease the effect of varying environmental
314 (meteorological) conditions between tests. In this trivariate model, we included sex and year
315 as fixed effects and ring number as a random effect. The phenotypic correlation between two
316 traits *A* and *B*, $r_{\mathrm{P_A,P_B}}$, and the between-individual correlation between the traits *A* and *B*,
317 $r_{\mathrm{ind_A,ind_B}}$, were estimated as follows (Snijders & Bosker 1999):

318 $$r_{\mathrm{P_A,P_B}} = \frac{\mathrm{Cov_{ind_A,ind_B}} + \mathrm{Cov_{\varepsilon_A,\varepsilon_B}}}{\sqrt{V_{\mathrm{P_A}} \times V_{\mathrm{P_B}}}} \quad (1)$$

319 $$r_{\mathrm{ind_A,ind_B}} = \frac{\mathrm{Cov_{ind_A,ind_B}}}{\sqrt{V_{\mathrm{ind_A}} \times V_{\mathrm{ind_B}}}} \quad (\text{eq. } 2)$$

where $\mathrm{Cov}_{\mathrm{ind_A,ind_B}}$ and $\mathrm{Cov}_{\varepsilon_A,\varepsilon_B}$ are the between- and the within-individual covariances between traits *A* and *B*, and $V_{\mathrm{P,\ A\ or\ B}}$ the total phenotypic variance (sum of the between- and within-individual variances) associated with trait A or B. Combining aggressiveness, boldness and neophobia score data for a given individual in a given year, we obtained 841 observations for which estimates for all three traits were available, and 152 observations for which only boldness and neophobia estimates were available. Among these 841 observations, 49 females and 52 males were repeatedly assessed over several years for all three traits, providing 100 and 111 repeated observations, respectively.

To estimate the additive genetic correlations between traits, we fitted a model with the same fixed effects but with the additive genetic effect instead of the individual ring as a random factor. Fitting both additive genetic and individual (permanent environment) effects together in a single model indeed led to convergence failures. We thus fitted only one effect at a time.

Implementation of Bayesian models

All statistical analyses were performed within the Bayesian framework in R v.3.3.2 (R Core Team, 2016). Both univariate and trivariate models were fitted using the function MCMCglmm ('MCMCglmm' R package, Hadfield, 2010). The pedigree was prepared using the function fixPedigree ('pedantics' R package, Morrissey & Wilson, 2010) and pruned using the function prunePed ('nadiv' R package, Wolak, 2012; see Appendix Table A1 for a description of the pedigree). Aggressiveness and boldness scores were fitted with a Poisson family (logit link), and neophobia scores with a Threshold family with the residual variance fixed to 10 (instead of the usual value of 1, to improve the mixing of the chains for low variances, which were expected from preliminary analyses; Hadfield, 2016). We used wide normally distributed priors for fixed effects (large variance V=10^8; Hadfield, 2016) and

parameter expanded χ^2 distributed priors with 1 degree of freedom for random effects. For the univariate models, we adjusted the number of iterations, burn-in and thinning interval for each model so as to obtain an effective sample size over 1500 (see Appendix) and autocorrelations of posterior samples below 0.1 in all cases. For the trivariate models, we used 4 x 10^6 iterations, a burn-in of 10^5 and a thinning interval of 2000 to reach the same criteria. We visually assessed the convergence of each MCMC chain and compared three chains per model using the Gelman & Rubin approach (gelman.diag and gelman.plot functions, 'coda' R package, Plummer et al., 2016). Following Nakagawa and Schielzeth (2010), we retained in our univariate models data from individuals tested only once. For all three traits, estimates are presented as posterior modes with the associated 95% credible intervals (CI). Variance, repeatability and heritability estimates are presented on the latent scale (R_{latent}, following Nakagawa & Schielzeth, 2010; h^2_{latent} following de Villemereuil, Schielzeth, Nakagawa, & Morrissey, 2016). We also provide repeatability and heritability estimates on the observed scale (R_{obs} using the QGicc function, h^2_{obs} using the QGparams function, from the 'QGglmm' R package; de Villemereuil, 2018; de Villemereuil, Schielzeth, Nakagawa, & Morrissey, 2015). It is not possible here to discuss all results on the observed scale, even though it is the scale of the realized behaviour, that is, where natural selection can act (see de Villemereuil et al., 2016). Indeed, the estimation of R_{obs} for ordinal traits (here the neophobia score) is complex and is not currently implemented in QGicc. To allow us to compare estimates between traits on the observed scale, we computed repeatabilities for neophobia using two alternative scores with nonordinal distributions, namely (1) the latency to enter the nestbox in the presence of the novel object for individuals that entered the nestbox (N=653, fitted with a Poisson distribution), and (2) whether the individual entered the nestbox in the presence of the novel object (N=1064, binary variable fitted with the threshold family). Moreover, the estimation of h^2_{obs} for ordinal traits provides one heritability estimate per level, which in the

370 case of an artificial categorization as here is not biologically relevant (de Villemereuil, 2018).
371 As the heritability estimates on the observed scale were fairly similar between neophobia
372 levels, we compared the range of heritability estimates found for neophobia with the
373 heritability estimates found on the observed scale for aggressiveness and boldness.
374 Correlations are provided on the latent scale.
375
376 ***Ethical note***
377 Permission for catching and ringing adults (here 838 yearlings, 1074 older birds) and
378 nestlings (here 9750) with individually numbered aluminium rings was granted every year by
379 the Ringing Centre from the Museum of Natural History in Stockholm (licence nb. 471: M015
380 to B.D.). Adults were caught in the nest, either directly (females during incubation) or using
381 swinging-door traps (both parents during nestling rearing). Traps were set for at most 30 - 60
382 min depending on nestling age (30 min when nestlings were 5 days old or younger), to avoid
383 nestling starvation if parents did not resume feeding during the catching period; traps were
384 checked every 5–10 min, and removed as soon as the adults had been caught. Catching
385 sessions started after 0600 hours to let birds feed and provision nestlings undisturbed for at
386 least 2 h after the night period (sunrise occurs at approximately 0400 hours during spring).
387 Adults were handled for 5 - 10 min and released straight after manipulation or (when catching
388 both parents during nestling feeding) kept until the partner was captured (up to 40 min
389 maximum). For nestling ringing, whole broods were taken directly from the nest and ringed
390 just beside the nestbox (for approximately 10 min); nestlings that were not handled were kept
391 warm using small heating packs. During the aggressiveness tests, we minimized disturbance
392 by approaching the nestbox as quietly as possible and hiding below a camouflage net.
393 Conversely, for the combined boldness/neophobia test, which aimed at measuring (or
394 controlling for) the reaction towards human presence, we on purpose approached the nestbox

conspicuously. During the boldness/neophobia test, nestling satiety was checked at the beginning and in between the two parts of the test, and the test was aborted if nestlings were begging too strongly, to avoid any harmful effect of temporarily decreased provisioning by parents. All the manipulations were performed in accordance with the Swedish legislation applying at the time.

RESULTS

Univariate models

The level of repeatability for aggressiveness was 0.18 on the latent scale and 0.03 on the observed scale (Table 1). In addition, aggressiveness score was repeatable within years (R_{latent}=0.22, 95% CI = [0.18; 0.26]; R_{obs}=0.04, 95% CI = [0.03; 0.06]) and between years (when averaging the aggressiveness score of 1 year; R_{latent}=0.26, 95% CI = [0.11; 0.38]; R_{obs}=0.11, 95% CI = [0.04; 0.20]). The level of repeatability for boldness was 0.11 on the latent scale and 0.10 on the observed scale (Table 1). Neophobia was slightly more repeatable that the other traits on the latent scale (Table 1) as well as on the observed scale when estimated from alternative nonordinal measures (R_{latent}=0.25, 95% CI = [0.06; 0.35] and R_{obs}=0.14, 95%CI = [0.04; 0.23] for the latency to enter the nestbox in the presence of the novel object; R_{latent}=0.44, 95% CI = [0.22; 0.67] and R_{obs}=0.29, 95%CI = [0.12; 0.44] for whether the individual entered the nestbox during the test or not). None of the three behavioural scores, however, were heritable (all 95% CI for V_A and h^2 values included zero; Table 1). Permanent environment effects explained 15% of the phenotypic variance in aggressiveness score (95% CI of V_{PE} = [0.28; 0.80]; Appendix Table A2) and parental identities (i.e. maternal and paternal identities summed) explained 25% of the phenotypic

variance in neophobia score (95% CI of $V_M + V_F$ = [0.38; 8.46], even though the lower limit of the 95% CI for each parent identity separately was 0: [0.00; 6.49] for V_M and [0.00; 5.75] for V_F). When excluding the maternal (paternal) identity from the model, the paternal (maternal) identity explained 19% (18%) of the phenotypic variance. Fitting the neophobia model without the maternal and paternal identities did not change the heritability estimate, revealing that these effects were not confounded with the additive genetic effect (results not detailed). Observer identity explained 7% of the phenotypic variance for aggressiveness (95% CI of $V_{Observer}$ = [0.14; 0.55]); paternal identity explained 11% of the phenotypic variance for aggressiveness but only when measuring aggressiveness as the first axis of a principal component analysis (see Appendix). All other variances were low (less than 4 % of the phenotypic variance) or not different from zero (Appendix Tables A1, A2 and A3).

Males were more aggressive than females, especially among yearlings (interaction sex*age, with yearling males as reference: 95% CI = [0.44, 0.81]; Appendix Fig. A1a). In addition, males were slightly shyer (longer latency to enter in the absence of a novel object) and less neophobic (shorter latency in the presence of a novel object) than females (with female as a reference 95% CI = [-0.12; -0.02] and [-2.78; -1.07] respectively; Appendix Fig. A1b), and this did not depend on age (see Appendix Tables A2 and A3 for sex*age interactions). In addition, individuals with larger broods were bolder (95% CI = [0.002; 0.06]; Table A3) and less neophobic (95% CI = [-1.03; -0.37]; Table A4). Regarding environmental effects, aggressiveness, boldness and neophobia scores depended on the year: individuals were less aggressive and less neophobic in 2011 than 2012 (Appendix Tables A2 and A4), and shyer in 2011 than 2013 (Appendix Table A3). Finally, individuals were more aggressive in the presence of their partner or neighbouring tits attracted by the stimulus, and during the first tests of the sequence (Appendix Table A2).

445 ***Trivariate model***

446 Aggressiveness and neophobia scores were phenotypically correlated: more aggressive
447 individuals were less neophobic (Table 2, Appendix Fig. A2a). Boldness and neophobia
448 scores were also phenotypically correlated: bolder individuals were less neophobic (Table 2,
449 Appendix Fig. A2b). The corresponding within-individual correlations were negative (Table
450 2). However, there was no phenotypic correlation between boldness and aggressiveness
451 (Table 2; see Appendix Table A5 for the full model output) and none of the between-
452 individual correlations differed from zero (Table 2). When accounting for an additive genetic
453 effect, none of the additive genetic covariances differed from zero (Table 2; see Table A6 for
454 the full model output).

455

456

457 **DISCUSSION**

458

459 In this study, we tested whether three personality traits (aggressiveness, boldness and
460 neophobia) had a genetic basis in a wild population of collared flycatchers and formed
461 (genetically based) behavioural syndromes during breeding, to shed light on constraints in
462 their possible evolution. None of the three traits were heritable and their repeatability
463 estimates were low (0.11–0.39 on the latent scale for all traits; 0.03–0.10 on the observed
464 scale for aggressiveness and boldness; 0.14 and 0.29 on the observed scale for nonordinal
465 measures of neophobia) compared to average estimates previously found for behavioural traits
466 in two meta-analyses (0.37 in Bell et al., 2009; 0.41 in Holtmann et al., 2017), suggesting
467 strong phenotypic plasticity in these traits. The repeatability originated mainly from
468 permanent environment effects for aggressiveness and from parental effects for neophobia. In
469 addition, we found that the three traits showed phenotypic covariance but no between-

individual covariance and no additive genetic covariance. The absence of behavioural syndromes among these personality traits may be due either to a lack of statistical power to detect between-individual covariances or to an absence of functional integration of these traits at the individual level and no genetic correlation at the population level.

Factors at the origin of behavioural trait repeatability

Our levels of repeatability, estimated both within and between years for aggressiveness score and between years for boldness and neophobia scores, were lower than usually reported for such behaviours: around 0.50 for aggressiveness and exploration, and around 0.40 for antipredator behaviours (Bell et al., 2009). Interestingly, the repeatability level estimated here for aggressiveness score was similar within and between years, contrary to the usual decrease observed when the time interval between recordings increases (Bell et al., 2009; Chervet, Zöttl, Schürch, Taborsky, & Heg, 2011; Dingemanse et al., 2012; Garamszegi et al., 2015; Holtmann et al., 2017; Wuerz & Krüger, 2015; but see David, Auclair, & Cézilly, 2012 for differences between traits). Overall, our lower levels of repeatability, especially on the observed scale, suggest higher plasticity, both within and between years compared to studies on other species.

The observed repeatability in aggressiveness score resulted mostly from permanent environment effects, which explained 15% of the phenotypic variance and 72% of the repeatability in aggressiveness score. Here, because we controlled for the identity of the parents, permanent environment effects could be linked to differences in individual condition or experience. Some measures of personality traits have indeed been found to depend on individual condition or experience (reviewed in Stamps & Groothuis, 2010). In our population, condition and experience also affect breeding habitat choice depending on social cues (e.g. Doligez, Danchin, Clobert, & Gustafsson, 1999; Doligez, Pärt, Danchin, Clobert, &

Gustafsson, 2004; Kivelä et al., 2014), which could shape individuals' response to the risk of competition for nest sites. Permanent environment effects may also include a dominance effect (Kruuk & Hadfield, 2007; Wilson et al., 2010), which could not be directly modelled here because full- and half-sib links were too rare in our pruned pedigree for running such complex models (Wilson et al., 2010).

In turn, the observed repeatability in neophobia resulted mostly from parental identities, which accounted for 25% of the phenotypic variance and 55% of the between-individual variance. Both pre- and postnatal parental effects have been found to affect exploration and neophobia behavioural responses later in life (e.g. nestling provisioning and exploration in birds, Carere, Drent, Koolhaas, & Groothuis, 2005; maternal hormones early in life and neophobia, Spencer & Verhulst, 2007; see the review in Groothuis & Maestripieri, 2013). However, parental identities did not explain between-individual differences in aggressiveness and boldness here, contrary to previous findings (e.g. Eising, Muller, & Groothuis, 2006; reviewed in Groothuis & Maestripieri, 2013). These behavioural responses may be more dependent on individual or local environmental conditions, in particular individual competitive ability and neighbour/predator presence or density, at the time of the test(s). To better understand how parental effects shape behavioural responses later in life in our study population, further experiments (e.g. nestling cross-fostering) would be necessary.

Personality traits with no genetic basis

We found no genetic basis for our three personality traits. A meta-analysis on personality traits in wild animal populations estimated an average heritability level of 0.28 for aggressiveness, 0.31 for boldness and 0.58 for exploration-avoidance (including estimates from novel environment and novel object tests; van Oers & Sinn, 2013). The absence of heritability for our personality traits here was not due to a lack of statistical power to detect

significant additive genetic variance based on our sample and social pedigree, because based on the same sample with the same pedigree, we obtained positive heritability estimates for tarsus and wing length (h^2=0.59, 95% CI = [0.44; 0.69] for tarsus and h^2=0.30, 95% CI = [0.11; 0.50] for wing length, while accounting for maternal and permanent environment effects) which are consistent with previous estimates in this population (h^2 = 0.53 and 0.51 for tarsus and wing length, respectively, in Merilä & Gustafsson, 1993). The absence of heritability in our personality traits was therefore likely to be the result of very low additive genetic variance combined with large environmental variance as illustrated, for instance, by between-year differences in behavioural scores, which reflected large variations in environmental conditions between the 3 years of our study (see Morinay, Forsman, Kivelä, Gustafsson, & Doligez, 2018 for differences between 2012 and 2013). Large environmental variance could originate from individuals being tested in different environments (including the social context) in different years, because between-year fidelity to the nestbox and/or partner is very low in this population (approximately 6.7% of 240 individuals bred in the same nestbox several years and 1.0% of 214 identified pairs were faithful over several years). This, however, limited the risk of pseudoreplication (Niemelä & Dingemanse, 2017).

Because we measured personality traits at the nest during breeding, the reaction of the partner may have affected the reaction of the focal bird during a behavioural test, as found here with a higher aggressiveness score when the partner was present. To account for this effect, we could have included the partner's identity and genetic background (i.e. pedigree) as random effects in our models (see Morinay et al., 2018 for an example in the same population; and Wolf, Brodie III, Cheverud, Moore, & Wade, 1998 for so-called indirect genetic effects). However, this could have led to pseudoreplication, because the behavioural score of the partner itself was most of the time also analysed in this dataset (e.g. around 74% of the females and 90% of the males had their partner tested). To keep exploring a response at

545 the individual level (rather than combining behavioural scores at the pair level), a solution
546 could be to fit a bivariate model of the two partners' responses and include both their
547 pedigree, permanent environment and parental effects in the model. This would, however,
548 require a larger data set than used here to reach sufficient statistical power to detect such
549 effects with such complex models. Furthermore, because the focal bird chooses at least partly
550 its partner (like its nest site), partner's effects can also be expected to be at least partly
551 included in the individual's genetic and permanent environment effects. Disentangling such
552 complex effects may require a more balanced sample of faithful and divorced pairs breeding
553 in the same and different sites over several years than observed in our population.
554
555 ***No personality syndrome?***
556 Phenotypic correlations were observed between our personality traits, even though they did
557 not constitute behavioural syndromes (i.e. there was no between-individual correlations): less
558 neophobic individuals were more aggressive and bolder. This was in line with previous
559 studies reporting bolder individuals to be more explorative in a novel environment (or less
560 neophobic in a novel object test; e.g. Garamszegi et al., 2009; van Oers, De Jong, Drent, &
561 van Noordwijk, 2004). Conversely, the absence of correlation between aggressiveness and
562 boldness partly contrasts with previous results reporting more aggressive individuals to be
563 bolder, as part of the proactive–reactive axis, in different species (Koolhaas et al., 1999; Sih,
564 Bell, Johnson, et al., 2004) including the collared flycatcher (Garamszegi et al., 2015).
565 The observed phenotypic correlations resulted solely from correlated changes in
566 behaviours between measurements for the same individuals, that is, within-individual
567 correlations. Within-individual correlations could be due to micro-environmental effects (e.g.
568 nestbox environment), to individual effects (e.g. long-term between-year plasticity but short-
569 term within-year behavioural constraints, for instance due to experience) or to correlated

570 measurement errors (Dingemanse & Dochtermann, 2013). Error correlation, however, is more
571 likely to occur between boldness and neophobia scores, which were extracted from the same
572 test and might both be correlated with feeding rate (with birds investing more in nestling
573 provisioning returning more rapidly to their nest in both situations), than between
574 aggressiveness and neophobia scores, which were measured several weeks apart by different
575 persons in different tests. To tease these sources of within-individual correlations apart,
576 aggressiveness, boldness and neophobia scores need to be estimated several times during the
577 same breeding season and possibly the same phase(s) of the reproductive cycle. The limited
578 number of individuals measured several times here (211 observations of 101 individuals),
579 however, is likely to be why we did not detect between-individual covariance (Dingemanse &
580 Dochtermann, 2013 recommended sample sizes of at least 200 individuals tested twice; see
581 also Garamszegi & Herczeg, 2012). Indeed, based on the same sample, we were not able to
582 obtain positive genetic or between-individual covariances between tarsus and wing length,
583 two morphological traits previously reported as genetically correlated in the same population
584 (Merilä & Gustafsson, 1993).

585 Even though our limited statistical power does not allow us to conclude the absence of
586 behavioural syndromes, this absence, if true, would suggest that selective pressures did not
587 yield or maintain a functional integration between the personality traits investigated here. A
588 true absence of behavioural syndromes among the traits we studied could be explained by
589 specific breeding conditions in our population, possibly altering the selective regimes
590 compared to other populations or species. In our population, the high availability of high-
591 quality nest sites (i.e. nestboxes, provided in excess since the early 1980s) may have released
592 joint selective pressures on exploration to find suitable nest sites and aggressiveness to
593 acquire and defend this resource against dominant competitors (in particular tit species) in a
594 natural context. In turn, providing nestboxes probably increased local breeding densities and

595 thereby competition for food resources during the nestling period, especially in a highly
596 synchronous species such as the collared flycatcher. Furthermore, our population is subjected
597 to very low nest predation rates, due to the absence of mustelid species on Gotland (Doligez
598 & Clobert, 2003), which may have released selective pressures on boldness through the
599 decrease in the need to defend the brood. Overall, these specific breeding conditions may
600 have strongly modified the selective regime for personality traits and for a functional
601 integration between them if they are costly.

602 In conclusion, we showed that aggressiveness, boldness and neophobia are repeatable
603 but not heritable traits and do not seem to form behavioural syndromes in our population of
604 collared flycatchers since only phenotypic correlations were observed between neophobia and
605 the other two traits. Our study thus brings insights on the evolutionary potential of these
606 personality traits alone and in interaction with each other during breeding in a wild population
607 experiencing particular breeding conditions (low competition for nest sites, low nest predation
608 rate). To understand the absence of individual covariance between, and heritable variations in,
609 personality traits in our population, a first step would be to investigate the fitness benefits (i.e.
610 reproductive success and survival) associated with each trait and their interactions. Plasticity
611 in the associations between personality traits should be selected for if the fitness costs and
612 benefits of expressing each trait relative to the others depend on the environmental (including
613 social) context (e.g. competition level or predation risk), which remains to be explored for
614 example by experimentally manipulating these environmental conditions.

615

616 **Conflict of interest**

617 We declare no conflict of interest.

618

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Appendix

Aggressiveness score

We estimated aggressiveness score using the following alternative measures and modelled them using the parameters given in parentheses. (1) Number of aggressive behaviours (movements < 2 m from the nestbox, stationary flights and attacks towards the decoys, chases of live birds) standardized per 15 min (Poisson family; number of iterations = 10^6; burn-in = 10^4; thinning interval = 400). (2) First axis of the PCA presented below (Table A7; Gaussian family; number of iterations = 10^6; burn-in = 10^4; thinning interval = 400). (3) Discrete score (threshold family, residual variance V_R = 10; number of iterations = 13 x 10^5; burn-in = 8 x 10^4; thinning interval = 700). This score was based on the distinction between activity (number of movements/min, including stationary flights and chases) performed far from (> 2 m) and close to (< 2 m) the nestbox, and on attacks, subdivided into six categories (Fig. A3a): **0:** individuals that performed no movements (either far from or close to the nestbox); 1: individuals that performed no attack or movements close to the nestbox and performed less than 0.440 movements/min far from (> 5 m) the nestbox; 2: individuals that performed no attack or movements close to the nestbox and performed more than 0.440 movements/min far from (> 5 m) the nestbox; 3: individuals that performed no attacks and less than 0.282 movements/min close to (< 2 m) the nestbox; 4: individuals that performed no attacks and between 0.282 and 0.784 movements/min close to (< 2 m) the nestbox; 5: individuals that performed no attacks and above 0.784 movements/min close to (< 2 m) the nestbox; 6:

individuals that performed attacks towards decoys(s). The thresholds were chosen so as to distribute individuals equally among categories for scores 1 and 2 on the one hand and scores 3, 4 and 5 on the other.

Boldness score

We estimated boldness score using the following alternative measures and modelled them using the parameters given in parentheses. (1) Maximum latency to enter the nestbox after human departure (from all individuals) minus the same latency for the focal individual on the focal test (Poisson family; number of iterations = 1 x 10^6; burn-in = 10^4; thinning interval = 500). (2) Inverse logarithmic ratio of the latency to enter the nestbox after human departure divided by the feeding rate for the remaining time after the first entrance in the nestbox. We divided by the feeding rate for this alternative measurement because the average feeding interval might have affected the latency to return to the nestbox, for example if individuals feeding more frequently entered the nestbox faster. The feeding rate after the first entrance was estimated as the average time interval between two feeding events by the focal birds, after the first entrance; it was thus only computable for individuals that fed at least twice in the period; for technical reasons, further individuals could not be used for this variable, leading to a final used data set of 641 observations (Gaussian family; number of iterations = 25 x 10^4; burn-in = 10^4; thinning interval = 100). (3) Discrete score (threshold family, V_R = 10; number of iterations = 3 x 10^6; burn-in = 2 x 10^5; thinning interval = 1000). The score was based on entrance in the nestbox during the first part of the boldness-neophobia test (no novel object) and latency to enter after human disturbance; individuals that did not enter were given a score of 0, and individuals that entered the nestbox were given a score of 1 - 5 based on five quantiles of the inverse latency to enter (Fig. A3b).

945

946 Neophobia score

947 We estimated neophobia score for individuals that entered the nestbox during the first period
948 of the test, using the following alternative measures and modelled them using the parameters
949 given in parentheses. (1) Discrete score based on the latency to enter the nestbox in the
950 presence of a novel object (second period of the test), discretized in four quantiles, the fifth
951 category including individuals that did not enter the nestbox at all in the presence of the novel
952 object (threshold family; V_R = 10; number of iterations = 15 x 10^5; burn-in = 2 x 10^5; thinning
953 interval = 500; Fig. A3c); (2) Binary variable separating individuals that did and did not enter
954 during the second period of the test (threshold family; V_R = 10; number of iterations = 10^6;
955 burn-in = 10^5; thinning interval = 500). (3) Latency to enter the nestbox in the presence of the
956 novel object, excluding the individuals that did not enter the nestbox during the second period
957 of the test (Poisson family; number of iterations= 15 x 10^5; burn-in = 15 x 10^4; thinning
958 interval = 500).
959 Results obtained for these alternative scores for the three behavioural traits are given in
960 Tables A2 to A4, and the main text presents the first score in each case.

961 **TABLES**

962

963 **Table 1.** Between-year repeatability and heritability estimates for aggressiveness, boldness
964 and neophobia scores

	R_{latent}	R_{obs}	h^2_{latent}	h^2_{obs}
Aggressiveness	0.18 *	0.03 *	0.00	0.00
	[0.15; 0.23]	[0.02; 0.04]	[0.00; 0.08]	[0.00; 0.008]
Boldness	0.11 *	0.10 *	0.00	0.00
	[0.01; 0.21]	[0.01; 0.19]	[0.00; 0.10]	[0.00; 0.09]
Neophobia	0.39 *		0.00	0.00 for all scores
	[0.25; 0.54]		[0.00; 0.15]	From [0.00; 0.00] to [0.00; 0.09]

965

966 Repeatabilities and heritabilities (posterior modes and 95% credible intervals) are given on
967 the latent scale (R_{latent}, h^2_{latent}) and on the observed scale (R_{obs} obtained with the QGicc function
968 and h^2_{obs} with the QGparams function from 'QGglmm' R package; de Villemereuil et al.
969 2016; de Villemereuil 2018). Asterisks indicate estimates whose 95% CI do not encompass
970 zero. For aggressiveness, estimates are given using all scores. For neophobia, we provide a
971 range of heritability values on the observed scale, since one value is provided per neophobia
972 score level (i.e. five values in total); however, we could not derive repeatability estimates on
973 the observed scale for ordinal variables.

974

975 **Table 2.** Phenotypic, between- and within-individual, and additive genetic correlations
976 between aggressiveness, boldness and neophobia scores
977

Correlation level	Aggressiveness - Neophobia	Boldness - Neophobia	Aggressiveness - Boldness
Phenotypic	-0.20*	-0.30*	0.02
	[-0.25; -0.12]	[-0.38; -0.23]	[-0.04; 0.10]
Between-individual	0.00	0.00	-0.02
	[-0.06; 0.06]	[-0.07; 0.05]	[-0.06; 0.06]
Within-individual	-0.28*	-0.43*	0.03
	[-0.38; -0.18]	[-0.53; -0.32]	[-0.04; 0.12]
Additive genetic	0.00	-0.01	0.01
	[-0.06; 0.006]	[-0.06; 0.06]	[-0.07; 0.06]

978
979 Posterior modes and 95% credible intervals (CI) on the latent scale are shown. Asterisks
980 indicate estimates whose 95% CI do not encompass zero.

Table A1. Detailed description of the collared flycatcher pedigree from the Gotland Island population

Pedigree statistics	
Records	2 218
Founders	1423
Maternities	728
Paternities	781
Mothers with ≥ 2 offspring	120
Fathers with ≥ 2 offspring	133
Full sibs	130
Maternal sibs	222
Maternal half sibs	92
Paternal sibs	230
Paternal half sibs	100
Maternal grandmothers	249
Maternal grandfathers	268
Paternal grandmothers	220
Paternal grandfathers	236
Maximum pedigree depth	15
Mean relatedness	5.29×10^{-4}

The pedigree statistics were obtained from all identified individuals involved in either aggressiveness, boldness or neophobia assays, and were extracted using the pedigreeStats and pedStatSummary functions from 'pedantics' R package (Morrissey & Wilson, 2010).

Table A2. Output of the univariate models fitting aggressiveness score

	Number of aggressive behaviours/15 min						PC1						D
	Without fixed effect				With fixed effects								
	Posterior mode	95% CI			Posterior mode	95% CI			Posterior mode	95% CI			P n
Fixed effects													
Intercept	1.22	[0.91	;	1.50] *	-0.29	[-0.72	;	0.26]	-0.98	[-1.30	;	-0.65] *	5
Sex (male)					0.63	[0.44	;	0.81] *	0.50	[0.35	;	0.62] *	1
Age (young)					0.23	[-0.01	;	0.40]	0.06	[-0.05	;	0.25]	0
Presence of tits (present)					0.34	[0.19	;	0.50] *	0.05	[-0.07	;	0.15]	0
Presence of flycatchers (present)					0.27	[-0.02	;	0.55]	0.29	[0.05	;	0.49] *	0
Presence of the partner (present)					0.62	[0.44	;	0.80] *	0.34	[0.20	;	0.44] *	1
Test number					-0.20	[-0.27	;	-0.14] *	-0.16	[-	;	-] *	-(

														0.20		0.11			
Dummy type (flycatcher)							0.08	[	-0.23	;	0.34	]		0.16	[	-0.01	;	0.38	]
Year (2012)							0.48	[	0.08	;	0.87	]	*	0.19	[	-0.05	;	0.50	]
Year (2013)							0.29	[	-0.09	;	0.67	]		0.10	[	-0.09	;	0.46	]
Sex*age (male*young)							0.28	[	0.03	;	0.61	]	*	0.24	[	0.09	;	0.50	]
Random effects																			
V_A	0.00	[	0.00	;	0.32	]		0.00	[	0.00	;	0.41	]		0.00	[	0.00	;	0.16	]
V_{PE}	0.60	[	0.28	;	0.80	]	*	0.52	[	0.11	;	0.71	]	*	0.19	[	0.03	;	0.28	]
V_M	0.00	[	0.00	;	0.20	]		0.00	[	0.00	;	0.17	]		0.00	[	0.00	;	0.12	]
V_F	0.00	[	0.00	;	0.20	]		0.00	[	0.00	;	0.24	]		0.23	[	0.06	;	0.37	]
V_{plot}	0.07	[	0.02	;	0.16	]	*	0.05	[	0.01	;	0.15	]	*	0.01	[	0.00	;	0.04	]
V_{decoy}	0.00	[	0.00	;	0.03	]		0.00	[	0.00	;	0.04	]		0.00	[	0.00	;	0.03	]
V_{song}	0.02	[	0.00	;	0.09	]		0.01	[	0.00	;	0.09	]		0.00	[	0.00	;	0.03	]

V_{obs}	0.25	[	0.14	;	0.55	]	*	0.30	[	0.15	;	0.63	]	*	0.06	[	0.02	;	0.16	]	*	0
$V\varepsilon$	2.56	[	0.91	;	1.50	]	*	2.11	[	1.97	;	2.34	]	*	1.49	[	1.40	;	1.57	]	*	1
Derived estimates																						
R_{latent}	0.18	[	0.15	;	0.23	]	*	0.18	[	0.13	;	0.22	]	*	0.23	[	0.17	;	0.27	]	*	0
R_{obs}	0.03	[	0.02	;	0.04	]	*	0.03	[	0.02	;	0.04	]	*	0.21	[	0.16	;	0.25	]	*	
h^2_{latent}	0.00	[	0.00	;	0.08	]		0.00	[	0.00	;	0.12	]		0.00	[	0.00	;	0.08	]		0
h^2_{obs}	0.00	[	0.00	;	0.01	]		0.00	[	0.00	;	0.01	]		0.00	[	0.00	;	0.07	]		P
$\overline{N}$		4680							3271							3271						
Effective sample size		> 2357							> 2234							> 2209						

Models for the general aggressiveness score, based on the number of aggressive behaviours standardized per 15 min, are shown without and with fixed effects. We also present the posterior modes and 95% credible intervals (CI) of models for alternative aggressiveness scores: the first axis of the principal component analysis (PC1) and the discrete score (see Appendix). Asterisks indicate estimates whose 95% CI do not encompass zero. For categorical fixed terms, estimates refer to the category indicated in parentheses. V_A, V_{PE}, V_M, V_F and V_ε refer to the additive genetic, permanent environment, maternal, paternal and residual variances, respectively. V_{plot}, V_{decoy}, V_{song}, V_{obs} refer to the variances associated with the plot, the decoy set used, the song track played and observer identity, respectively. N is the sample size of the data set used in the model.

Repeatability and heritability estimates are given both on the latent scale (R_{latent}, h^2_{latent}) and on the observed scale (R_{obs}, h^2_{obs}) whenever these could be estimated.

11 **Table A3 Output of the models fitting boldness scores.**

	Maximum latency–individual latency						Latency/feeding rate				Discrete sc
	Without fixed effects			Without fixed effects							
	Posterior mode	95% CI		Posterior mode	95% CI		Posterior mode	95% CI			Posterior mode
Fixed effects											
Intercept	8.00	[7.97	; 8.04] *	8.05	[7.96	; 8.10] *	-3.78	[-4.07	; -3.46] *		4.80
Sex (male)				-0.04	[-0.12	; 0.02]	-0.10	[-0.29	; 0.13]		-1.33
Age (young)				-0.02	[-0.10	; 0.05]	-0.12	[-0.39	; 0.13]		-0.65
Year (2012)				-0.03	[-0.10	; 0.04]	0.05	[-0.19	; 0.43]		-0.38
Year (2013)				0.08	[0.00	; 0.16] *	0.47	[0.10	; 0.73] *		1.46
No. of chicks				0.03	[0.00	; 0.06] *	0.26	[0.15	; 0.33] *		0.36
Sex*age (male*young)				0.03	[-0.08	; 0.13]	0.30	[-0.18	; 0.55]		0.60
Random effects											
V_A	0.00	[0.00	; 0.02]	0.00	[0.00	; 0.03]	0.00	[0.00	; 0.20]		0.01

V_{PE}	0.00	[	0.00	;	0.02	]	0.00	[	0.00	;	0.03	]	0.00	[	0.00	;	0.25	]	0.04
V_{M}	0.00	[	0.00	;	0.02	]	0.00	[	0.00	;	0.02	]	0.00	[	0.00	;	0.13	]	0.01
V_{F}	0.00	[	0.00	;	0.01	]	0.00	[	0.00	;	0.01	]	0.00	[	0.00	;	0.15	]	0.01
V_{plot}	0.00	[	0.00	;	0.00	]	0.00	[	0.00	;	0.01	]	0.00	[	0.00	;	0.07	]	0.00
V_{obs}	0.00	[	0.00	;	0.00	]	0.00	[	0.00	;	0.00	]	0.00	[	0.00	;	0.03	]	0.00
$V\varepsilon$	0.16	[	0.14	;	0.17	] *	0.13	[	0.11	;	0.16	] *	1.20	[	0.93	;	1.37	] *	10.00
Derived estimates																			
R_{latent}	0.11	[	0.01	;	0.21	] *	0.19	[	0.07	;	0.31	] *	0.12	[	0.01	;	0.29	] *	0.28
R_{obs}	0.10	[	0.01	;	0.19	] *	0.12	[	0.02	;	0.23	] *	0.04	[	0.01	;	0.14	] *	
h^2_{latent}	0.00	[	0.00	;	0.10	]	0.00	[	0.00	;	0.18	]	0.00	[	0.00	;	0.14	]	0.00
h^2_{obs}	0.00	[	0.00	;	0.09	]	0.00	[	0.00	;	0.09	]	0.00	[	0.00	;	0.06	]	Posterior m
N			1064						914						641				
Effective sample size			> 1718						> 1979						> 2053				

Models for the general boldness estimate (maximum latency to enter the nestbox after human disturbance observed in the entire data set minus the individual latency) are shown without and with fixed effects. We also present the posterior modes and 95% credible intervals (CI) of models

for alternative boldness estimates: the log-transformed and inverse ratio of the latency to enter the nestbox after human disturbance over the feeding rate during the time remaining and the discrete score based on the latency (see Appendix). Asterisks indicate estimates whose 95% CI do not encompass zero. For categorical fixed terms, estimates refer to the category indicated in parentheses. V_A, V_{PE}, V_M, V_F and V_ε refer to the additive genetic, permanent environment, maternal, paternal and residual variances, respectively. V_{plot} and V_{obs} refer to the variances associated with the plot and the observer identity, respectively. N is the sample size of the data set used in the model. Repeatability and heritability estimates are given both on the latent scale (R_{latent}, h^2_{latent}) and on the observed scale (R_{obs}, h^2_{obs}) whenever these could be estimated. Repeatability and heritability estimates for the latency/feeding rate model (fitted with a Gaussian distribution) differ between the latent and observed scales because we accounted for the inverse log transformation in the QGicc and QGparams functions ('QGglmm' R pachage; de Villemereuil et al. 2016).

22 **Table A4. Output of the models fitting neophobia scores**

	Discrete score						Binary			La
	Without fixed effect			With fixed effects						
	Posterior mode	95% CI		Posterior mode	95% CI		Posterior mode	95% CI		Po m
Fixed effects										
Intercept	4.22	[3.40	; 4.94] *	5.08	[3.63	; 6.38] *	-1.67	[-3.39	; -0.30] *	6.
Sex (male)				-1.89	[-2.78	; -1.07] *	-1.45	[-2.44	; -0.41] *	-0
Age (young)				-0.18	[-1.01	; 0.76]	0.07	[-1.26	; 0.92]	-0
Year (2012)				1.62	[0.72	; 2.93] *	2.19	[0.59	; 3.85] *	0.
Year (2013)				1.21	[-0.05	; 2.37]	1.84	[-0.06	; 3.34]	0.
Boldness				-1.58	[-1.93	; -1.12] *	-1.34	[-1.95	; -0.93] *	-0

No. of chicks						-0.71	[	-1.03	;	-0.37	]	*	-0.71	[	-1.15	;	-0.30	]	*	-0
Sex*age (male*young)						-1.02	[	-2.44	;	0.11	]		-0.79	[	-2.75	;	0.69	]		-0
Random effects																				
V_A	0.02	[	0.00	;	2.90	]	0.03	[	0.00	;	5.76	]		0.08	[	0.00	;	8.41	]		0.
V_{PE}	0.04	[	0.00	;	5.77	]	0.06	[	0.00	;	9.03	]		0.05	[	0.00	;	10.70	]		0.
V_M	0.04	[	0.00	;	6.49	]	0.06	[	0.00	;	8.81	]		0.12	[	0.00	;	22.53	]		0.
V_F	0.02	[	0.00	;	5.75	]	0.03	[	0.00	;	7.85	]		0.04	[	0.00	;	11.33	]		0.
V_{plot}	0.98	[	0.26	;	2.50	] *	1.01	[	0.27	;	3.42	]	*	0.80	[	0.00	;	3.08	]		0.
V_{obs}	0.28	[	0.02	;	1.01	] *	0.38	[	0.00	;	1.41	]		0.80	[	0.14	;	3.14	]	*	0.
$V\varepsilon$	10.00	[	10.00	;	10.00	]	10.00	[	10.00	;	10.00	]		10.00	[	10.00	;	10.00	]		0.
Derived estimates																				
R_{latent}	0.39	[	0.25	;	0.54	] *	0.50	[	0.33	;	0.63	]	*	0.56	[	0.35	;	0.75	]	*	0.
R_{obs}														0.29	[	0.17	;	0.48	]	*	0.
h^2_{latent}	0.00	[	0.00	;	0.15	]	0.00	[	0.00	;	0.22	]		0.00	[	0.00	;	0.26	]		0.

h^2_{obs}	Posterior mode range: 0.00-0.00	Posterior mode range: 0.00-0.00	0.00	[0.00 ; 0.14]	0.
N	1 064	914	914		
Effective sample size	> 2302	> 2317	> 1583		

Models for the general neophobia discrete score, based on the latency to enter the nestbox in presence of the novel object, are shown without and with fixed effects We also present the posterior modes and 95% credible intervals (CI) of models for alternative neophobia estimates: a binary variable (individual entered versus did not enter in the presence of the novel object) and a continuous latency to enter the nestbox, for individuals that entered in the presence of the novel object (see Appendix). Asterisks indicate estimates whose 95% CI do not encompass zero. For categorical fixed terms, estimates refer to the category indicated in parentheses. V_A, V_{PE}, V_M and V_F refer to the additive genetic, permanent environment, maternal and paternal variances, respectively. The residual variance V_ε was set to 10. V_{plot} and V_{obs} refer to the variances associated with the plot and the observer identity, respectively. *N* is the sample size of the data set used in the model. Repeatability and heritability estimates are given both on the latent scale (R_{latent}, h^2_{latent}) and the observed scale (R_{obs}, h^2_{obs}) whenever these could be estimated.

Table A5. Output of the trivariate model fitting aggressiveness, boldness and neophobia scores with an individual random effect

	Posterior mode	95% CI				
Fixed effects						
Aggressiveness	1.73	[	1.65	;	1.82	] *
Boldness	7.97	[	7.93	;	8.02	] *
Neophobia	4.44	[	3.79	;	5.00	] *
Sex (male)	-0.02	[	-0.06	;	0.04	]
Year (2012)	0.04	[	-0.02	;	0.09	]
Year (2013)	0.09	[	0.03	;	0.15	] *
Random effects						
$V_{\text{ind, aggressiveness}}$	0.42	[	0.16	;	0.66	] *
$V_{\text{ind, boldness}}$	0.02	[	0.00	;	0.04	]
$V_{\text{ind, neophobia}}$	6.14	[	2.96	;	10.27	] *
$V_{\varepsilon\text{, aggressiveness}}$	1.36	[	1.09	;	1.62	] *

$V_{\varepsilon,\ \text{boldness}}$	0.13	[	0.12	;	0.16	]	*
$V_{\varepsilon,\ \text{neophobia}}$	10.00	[	10.00	;	10.00	]	
$\text{Cov}_{\text{ind, aggressiveness-neophobia}}$	0.00	[	-0.09	;	0.11	]	
$\text{Cov}_{\text{ind, aggressiveness-boldness}}$	0.00	[	-0.01	;	0.01	]	
$\text{Cov}_{\text{ind, boldness-neophobia}}$	0.00	[	-0.03	;	0.02	]	
$\text{Cov}_{\varepsilon,\ \text{aggressiveness-neophobia}}$	-1.03	[	-1.37	;	-0.63	]	*
$\text{Cov}_{\varepsilon,\ \text{aggressiveness-boldness}}$	0.02	[	-0.02	;	0.05	]	
$\text{Cov}_{\varepsilon,\ \text{boldness-neophobia}}$	-0.49	[	-0.62	;	-0.36	]	*
N		1689					
Effective sample size:		>1740					

V stands for variance terms and Cov for covariance terms (posterior mode and 95% credible interval (CI)). The individual effect was the only random term included; 'ind' and 'ε' stand for between-individual and residual terms, respectively. The residual variance for neophobia score was fixed to 10 (see text for the distributions used for the three scores). Asterisks indicate estimates whose 95% CI do not encompass zero.

Table A6. Output of the trivariate model fitting aggressiveness, boldness and neophobia scores with an additive genetic effect

	Posterior mode	95% CI				
Fixed effects						
Aggressiveness	1.71	[	1.62	;	1.79	] *
Boldness	7.97	[	7.93	;	8.02	] *
Neophobia	4.12	[	3.55	;	4.62	] *
Sex (male)	-0.01	[	-0.05	;	0.04	]
Year (2012)	0.04	[	-0.02	;	0.09	]
Year(2013)	0.09	[	0.03	;	0.15	] *
Random effects						
$V_{A,\ aggressiveness}$	0.15	[	0.00	;	0.38	]
$V_{A,\ boldness}$	0.02	[	0.00	;	0.04	]
$V_{A,\ neophobia}$	4.81	[	2.07	;	7.52	] *
$V_{\varepsilon,\ aggressiveness}$	1.46	[	1.29	;	1.78	] *

$V_{\varepsilon,\ \text{boldness}}$	0.14	[	0.12	;	0.16	]	*
$V_{\varepsilon,\ \text{neophobia}}$	10.00	[	10.00	;	10.00	]	
$\text{Cov}_{\text{A, aggressiveness-neophobia}}$	0.00	[	-0.06	;	0.07	]	
$\text{Cov}_{\text{A, aggressiveness-boldness}}$	0.00	[	0.00	;	0.00	]	
$\text{Cov}_{\text{A, boldness-neophobia}}$	0.00	[	-0.02	;	0.02	]	
$\text{Cov}_{\varepsilon,\ \text{aggressiveness-neophobia}}$	-1.01	[	-1.35	;	-0.64	]	*
$\text{Cov}_{\varepsilon,\ \text{aggressiveness-boldness}}$	0.01	[	-0.02	;	0.05	]	
$\text{Cov}_{\varepsilon,\ \text{boldness-neophobia}}$	-0.48	[	-0.61	;	-0.37	]	*
N		1686					
Effective sample size		> 1816					

V stands for variance terms and Cov for covariance terms (posterior mode and 95% credible interval (CI)). The additive genetic effect was the only random term included; 'A' and 'ε' stand for additive genetic and residual terms, respectively. The residual variance for neophobia score was fixed to 10 (see text for the distributions used for the three scores). Asterisks indicate estimates whose 95% CI do not encompass zero. Positive additive genetic variances are found here because permanent environment and parental effects are not taken into account.

Table A7. Output of a principal component analysis of behaviours recorded during the aggressiveness assays

	Coordinates		Contribution	
	PC1	PC2	PC1	PC2
Moves < 2 m from the nestbox	0.83	-0.12	36.12	0.95
Stationary flights towards the decoy	0.73	-0.21	27.74	2.99
Attacks towards the decoy	0.67	-0.24	23.24	4.06
Moves between 2 and 5 m from the nestbox	0.42	0.65	9.40	29.96
Moves between 5 and 10 m from the nestbox	-0.03	0.74	0.03	38.49
Chases of live birds	0.26	0.58	3.47	23.55

We used the function PCA from the 'FactoMineR' R package (Lê, Josse, & Husson, 2008).

FIGURE CAPTIONS

Figure A1. Sex differences in aggressiveness, boldness and neophobia scores (means ± 95% confidence interval). (a) Between-sexes differences in average aggressiveness score for a given individual in a given year depending on age (yearling versus older). (b) Between-sexes differences in the latency to return after human disturbance for the period without a novel object, as a proxy of (inverse) boldness, and for the period with the novel object, as a proxy of neophobia. Number of observations is indicated near each estimate.

Figure A2. (a) Aggressiveness and (b) boldness scores depending on the neophobia score (means ± 95% confidence interval). See text for the definitions of the scores. Aggressiveness is here the averaged value of all scores for a given individual in a given year. Number of observations is indicated near each estimate.

Figure A3. Distribution of the discrete scores for (a) aggressiveness, (b) boldness and (c) neophobia. Number of observations is indicated above each bar.

22 **Figure A1**

23

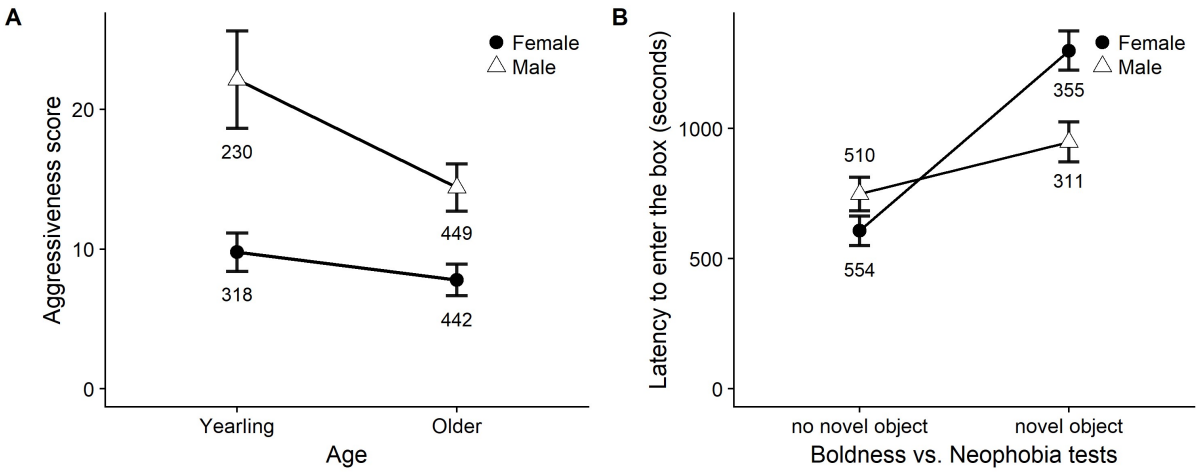
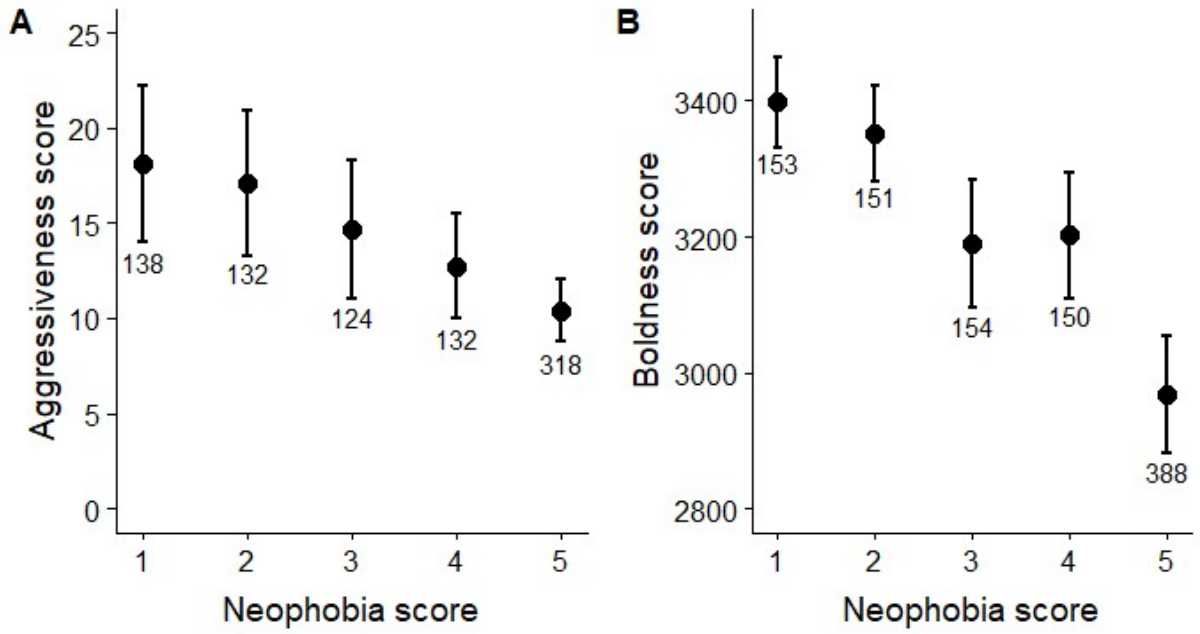
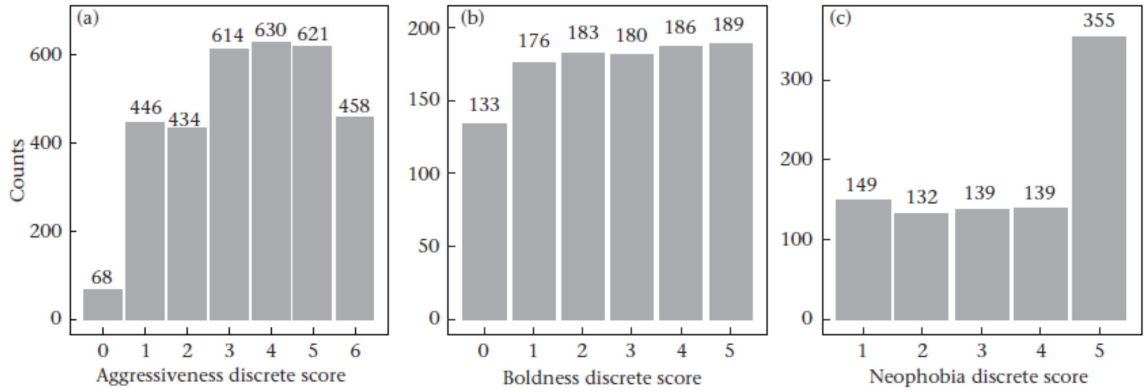


Figure A2


A
Aggressiveness score
25
20
15
10
5
0
138
132
124
132
318
1
2
3
4
5
Neophobia score
B
Boldness score
3400
3200
3000
2800
153
151
154
150
388
1
2
3
4
5
Neophobia score

26 **Figure A3**


(a)
Counts
600
400
200
0
68
446
434
614
630
621
458
0 1 2 3 4 5 6
Aggressiveness discrete score
(b)
200
150
100
50
0
133
176
183
180
186
189
0 1 2 3 4 5
Boldness discrete score
(c)
300
200
100
0
149
132
139
139
355
1 2 3 4 5
Neophobia discrete score