Introduction

Conclusion

MASSICCC: A SaaS Platform for Clustering and Co-Clustering of Mixed Data

https://massiccc.lille.inria.fr/

C. Biernacki

with B. Auder, G. Celeux, J. Demont, F. Langrognet, V. Kubicki, C. Poli, J. Renault, S. lovleff

APSEM2019 : éco-systèmes pour la science ouverte et recherche par les données 15-18 octobre 2019, ENSEEIHT, Toulouse







Take home message

- **Result accuracy** is impacted by the retained model for variety, veracity, d, n...
- Velocity is impacted by human handmade pre/post-processing



MASSICCC reduces pre- and post-processing by using models

Conclusion



1 Introduction

- 2 Model-based clustering
- 3 Mixmod in MASSICCC
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCC
- 6 Conclusion

Conclusion

MASSICCC?

massiccc.lille.inria.fr



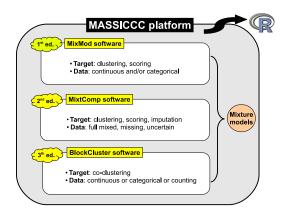
Developed by Inria

RY IT I

SaaS: Software as a Service

Conclusion

MASSICCC??



A high quality and easy to use web platform where are transfered mature research clustering (and more) software towards (non academic) professionals Introduction

BlockCluster

Conclusion

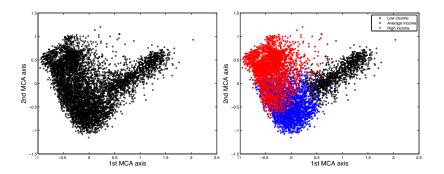
Here is the computer you need!



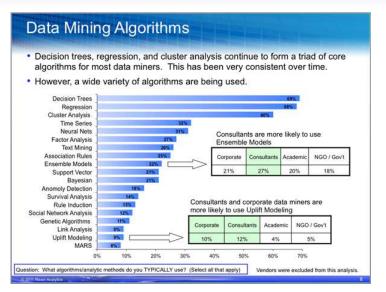




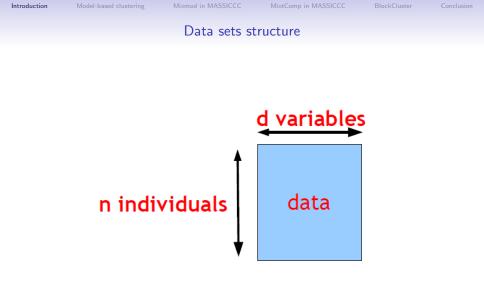
Detect hidden structures in data sets



Clustering everywhere¹

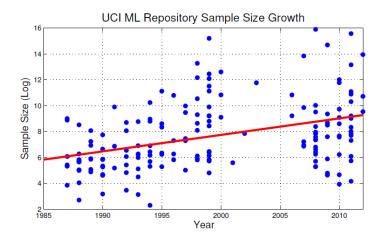


¹Rexer Analytics's Annual Data Miner Survey is the largest survey of data mining, data science, and analytics professionals in the industry (survey of 2011)



8/79

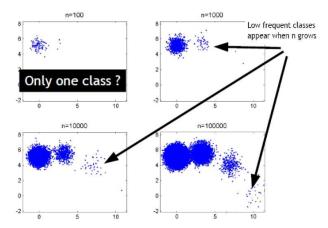
Large data sets²



 $^2 {\rm S.}$ Alelyani, J. Tang and H. Liu (2013). Feature Selection for Clustering: A Review. Data Clustering: Algorithms and Applications, ${\bf 29}$

9/79

An opportunity for detecting weak signal

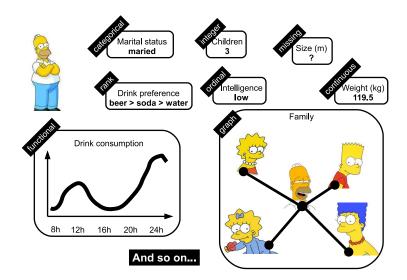


Introduction

BlockCluster

Conclusion

Todays features: full mixed/missing



Notations

- **Data:** *n* individuals: $\mathbf{x} = {\mathbf{x}_1, \dots, \mathbf{x}_n} = {\mathbf{x}^O, \mathbf{x}^M}$ in a space \mathcal{X} of dimension *d*
 - Observed individuals x^O
 Missing individuals x^M
- Aim: estimation of the partition \mathbf{z} and the number of clusters KPartition in K clusters G_1, \ldots, G_K : $\mathbf{z} = (\mathbf{z}_1, \ldots, \mathbf{z}_n), \mathbf{z}_i = (z_{i1}, \ldots, z_{iK})'$

$$\mathbf{x}_i \in G_k \quad \Leftrightarrow \quad z_{ih} = \mathbb{I}_{\{h=k\}}$$

	Pa	rtitio	n z	\Leftrightarrow	Group			
?	0.5	red	5	?	?	?	\Leftrightarrow	???
0.3	0.1	green	3	?	?	?	\Leftrightarrow	???
0.3	0.6	{red,green}	3	?	?	?	\Leftrightarrow	???
0.9	[0.25 0.45]	red	?	?	?	?	\Leftrightarrow	???
\downarrow	· ↓ ·	\downarrow	\downarrow					
continuous	continuous	categorical	integer					

Conclusion





- 2 Model-based clustering
- **3** Mixmod in MASSICCC
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCC
- 6 Conclusion

Conclusion

Parametric mixture model

Parametric assumption:

$$\mathsf{p}_k(\mathsf{x}_1) = \mathsf{p}(\mathsf{x}_1; \boldsymbol{\alpha}_k)$$

thus

$$\mathsf{p}(\mathsf{x}_1) = \mathsf{p}(\mathsf{x}_1; \boldsymbol{\theta}) = \sum_{k=1}^{K} \pi_k \mathsf{p}(\mathsf{x}_1; \boldsymbol{\alpha}_k)$$

Mixture parameter:

$$oldsymbol{ heta} = (oldsymbol{\pi}, oldsymbol{lpha})$$
 with $oldsymbol{lpha} = (oldsymbol{lpha}_1, \dots, oldsymbol{lpha}_K)$

Model: it includes both the family $p(\cdot; \alpha_k)$ and the number of groups K

 $\mathbf{m} = \{ \mathsf{p}(\mathbf{x}_1; \boldsymbol{\theta}) : \boldsymbol{\theta} \in \Theta \}$

The number of free continuous parameters is given by

 $\nu = \dim(\Theta)$

Clustering becomes a well-posed problem...

Conclusion

The clustering process in mixtures

1 Estimation of θ by $\hat{\theta}$

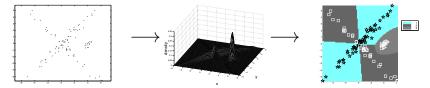
2 Estimation of the conditional probability that $\mathbf{x}_i \in G_k$

$$t_{ik}(\hat{\theta}) = p(Z_{ik} = 1 | \mathbf{X}_i = \mathbf{x}_i; \hat{\theta}) = \frac{\hat{\pi}_k p(\mathbf{x}_i; \hat{\alpha}_k)}{p(\mathbf{x}_i; \hat{\theta})}$$

3 Estimation of z_i by maximum a posteriori (MAP)

$$\hat{z}_{ik} = \mathbb{I}_{\{k = \arg \max_{h=1,...,K} t_{ih}(\hat{\theta})\}}$$

Model selection: BIC, ICL, ...

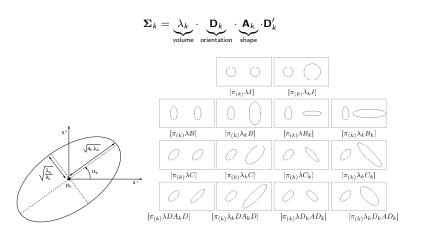






- 2 Model-based clustering
- 3 Mixmod in MASSICCC
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCC
- 6 Conclusion

Only continuous features: 14 models on Σ_k



Only categorical variables: latent class model

Categorical variables: d variables with m_j modalities each, $\mathbf{x}_j^j \in \{0,1\}^{m_j}$ and

 $\mathbf{x}_i^{jh} = 1 \quad \Leftrightarrow \quad \text{variable } j \text{ of } \mathbf{x}_i \text{ takes level } h$

Conditional independence:

$$\mathsf{p}(\mathbf{x}_i; \boldsymbol{\alpha}_k) = \prod_{j=1}^d \prod_{h=1}^{m_j} (\alpha_k^{jh})^{x_i^{jh}}$$

and

$$\alpha_k^{jh} = \mathsf{p}(\mathbf{x}_i^{jh} = 1 | z_{ik} = 1)$$

with $\boldsymbol{\alpha}_k = (\alpha_k^{jh}; j = 1, \dots, d; h = 1, \dots, m_j)$

Mixing continuous and categorical data: full local independence

Combine continuous and categorical data

$$\mathbf{x}_1 = (\mathbf{x}_1^{cont}, \mathbf{x}_1^{cat})$$

The proposed solution is to mixed both types by inter-type conditional independence

$$\mathsf{p}(\mathsf{x}_1; \boldsymbol{\alpha}_k) = \mathsf{p}(\mathsf{x}_1^{cont}; \boldsymbol{\alpha}_k^{cont}) \times \mathsf{p}(\mathsf{x}_1^{cat}; \boldsymbol{\alpha}_k^{cat})$$

In addition, for symmetry between types, intra-type conditional independence

Only need to define the univariate pdf for each variable type!

Continuous: Gaussian

Categorical: multinomial

Estimation of θ by *complete*-likelihood

Maximize the complete-likelihood over (θ, z)

$$\ell_{c}(\boldsymbol{\theta}; \mathbf{x}, \mathbf{z}) = \sum_{i=1}^{n} \sum_{k=1}^{K} z_{ik} \ln \{\pi_{k} \mathsf{p}(\mathbf{x}_{i}; \boldsymbol{\alpha}_{k})\}$$

Equivalent to traditional methods

Metric	M = I	M free	\mathbf{M}_k free
Gaussian model	$[\pi\lambda I]$	$[\pi\lambda C]$	$[\pi \lambda_k C_k]$

- Bias of $\hat{\theta}$: heavy if poor separated clusters
- Associated optimization algorithm: CEM (see later)
- **CEM** with $[\pi\lambda I]$ is strictly equivalent to K-means
- CEM is simple et fast (convergence with few iterations)

Estimation of θ by *observe*-likelihood

Maximize the observe-likelihood on θ

$$\ell(\boldsymbol{ heta}; \mathbf{x}) = \sum_{i=1}^{n} \ln p(\mathbf{x}_i; \boldsymbol{ heta})$$

- **Convergence** of $\hat{\theta}$, asymptotic efficiency, asymptotically unbiased
- General algorithm for missing data: EM
- EM is simple but slower than CEM
- Interpretation: it is a kind of fuzzy clustering



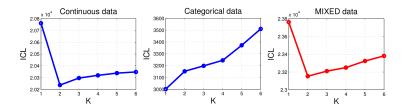
- Initialization: θ^0
- Iteration n^oq:
 - Step E: estimate probabilities $\mathbf{t}^q = \{t_{ik}(\boldsymbol{\theta}^q)\}$
 - Step C: classify by setting $\mathbf{t}^q = MAP(\{t_{ik}(\boldsymbol{\theta}^q)\})$
 - Step M: maximize $\theta^{q+1} = \arg \max_{\theta} \ell_c(\theta; \mathbf{x}, \mathbf{t}^q)$
- Stopping rule: iteration number or criterion stability

Properties

- $\blacksquare \oplus$: simplicity, monotony, low memory requirement
- \ominus : local maxima (depends on θ^0), linear convergence (EM)

Prostate cancer data (without mixing data)

- Individuals: n = 475 patients with prostatic cancer grouped on clinical criteria into two Stages 3 and 4 of the disease
- Variables: *d* = 12 pre-trial variates were measured on each patient, composed by eight continuous variables (age, weight, systolic blood pressure, diastolic blood pressure, serum haemoglobin, size of primary tumour, index of tumour stage and histolic grade, serum prostatic acid phosphatase) and four categorical variables with various numbers of levels (performance rating, cardiovascular disease history, electrocardiogram code, bone metastases)

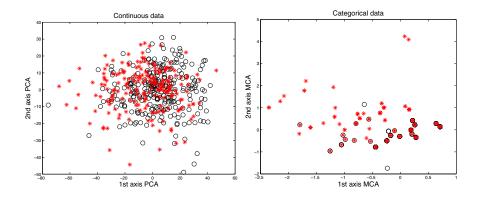


Model: cond. indep. $p(\mathbf{x}_1; \boldsymbol{\alpha}_k) = p(\mathbf{x}_1; \boldsymbol{\alpha}_k^{cont}) \cdot p(\mathbf{x}_1; \boldsymbol{\alpha}_k^{cat})$

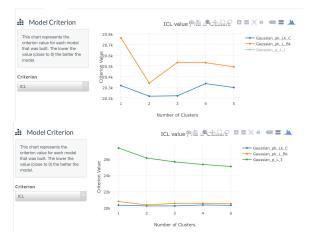
Conclusion

Prostate cancer data (without missing data)

Variables	Conti	nuous	Categ	gorical	Mi	×ed
Error (%)	9.	46	47	.16	8.	63
True \setminus estimated group	1	2	1	2	1	2
Stage 3	247	26	142	131	252	21
Stage 4	19	183	120	82	20	182

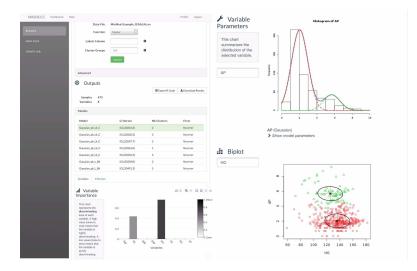


Continuous data

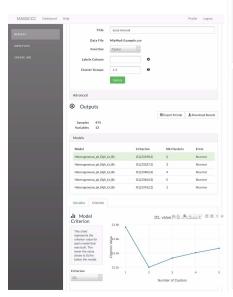


Conclusion

Continuous data



Mixed data



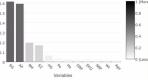
Model	Criterion	Nb Clusters	Error
Heterogeneous_pk_Ekjh_Lk_Bk	ICL(23198.3)	2	No error
Heterogeneous_pk_Ekjh_Lk_Bk	ICL(23327.2)	3	No error
Heterogeneous_pk_Ekjh_Lk_Bk	ICL(23402.6)	4	No error
Heterogeneous_pk_Ekjh_Lk_Bk	ICL(23464.3)	5	No error
Heterogeneous_pk_Ekjh_Lk_Bk	ICL(23762.2)	1	No error

Variables Criterion











AP

Y Hide model parameters Class 2

mean: 3.9268, sigma: 2.8781 mean: 1.6476, sigma: 0.2509





- 2 Model-based clustering
- 3 Mixmod in MASSICCC
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCC



Full mixed data: conditional independence everywhere³

The aim is to combine continuous, categorical, integer data, ordinal, ranking and functional data

$$\mathbf{x}_1 = (\mathbf{x}_1^{cont}, \mathbf{x}_1^{cat}, \mathbf{x}_1^{int}, \ldots)$$

The proposed solution is to mixed all types by inter-type conditional independence

$$\mathsf{p}(\mathsf{x}_1; \boldsymbol{\alpha}_k) = \mathsf{p}(\mathsf{x}_1^{cont}; \boldsymbol{\alpha}_k^{cont}) \times \mathsf{p}(\mathsf{x}_1^{cat}; \boldsymbol{\alpha}_k^{cat}) \times \mathsf{p}(\mathsf{x}_1^{int}; \boldsymbol{\alpha}_k^{int}) \times \dots$$

In addition, for symmetry between types, intra-type conditional independence

Only need to define the univariate pdf for each variable type!

- Continuous: Gaussian
- Categorical: multinomial
- Integer: Poisson
- • •

³MixtComp software on the MASSICCC platform: https://massiccc.lille.inria.fr/

Missing data: MAR assumption and estimation

Assumption on the missingness mecanism

Missing At Randon (MAR): the probability that a variable is missing does not depend on its own value given the observed variables.

Observed log-likelihood...

$$\ell(\boldsymbol{\theta}; \mathbf{x}^{\mathrm{O}}) = \sum_{i=1}^{n} \log \left(\sum_{k=1}^{K} \pi_{k} \mathbf{p}(\mathbf{x}_{i}^{\mathrm{O}}; \boldsymbol{\alpha}_{k}) \right) = \ln \left[\sum_{k=1}^{K} \pi_{k} \underbrace{\int_{\mathbf{x}_{i}^{M}} \mathbf{p}(\mathbf{x}_{i}^{O}, \mathbf{x}_{i}^{M}; \boldsymbol{\alpha}_{k}) d\mathbf{x}_{i}^{M}}_{\text{MAR assumption}} \right]$$

SEM algorithm⁴

A SEM algorithm to estimate heta by maximizing the observed-data log-likelihood

- Initialisation: $\theta^{(0)}$
- Iteration nb q:
 - **E**-step: compute conditional probabilities $p(\mathbf{x}^{M}, \mathbf{z} | \mathbf{x}^{0}; \boldsymbol{\theta}^{(q)})$
 - S-step: draw $(x^{M(q)}, z^{(q)})$ from $p(x^M, z|x^0; \theta^{(q)})$
 - M-step: maximize $\theta^{(q+1)} = \arg \max_{\theta} \ln p(\mathbf{x}^{O}, \mathbf{x}^{M(q)}, \mathbf{z}^{(q)}; \theta)$
- Stopping rule: iteration number

Properties: simpler than EM and interesting properties!

- Avoid possibly difficult E-step in an EM
- Classical M steps
- Avoids local maxima
- \blacksquare The mean of the sequence $(heta^{(q)})$ approximates $\hat{ heta}$
- The variance of the sequence $(\theta^{(q)})$ gives confidence intervals

⁴MixtComp software on the MASSICCC platform: https://massiccc.lille.inria.fr/

Prostate cancer data (with missing data)⁵

- Individuals: 506 patients with prostatic cancer grouped on clinical criteria into two Stages 3 and 4 of the disease
- Variables: d = 12 pre-trial variates were measured on each patient, composed by eight continuous variables (age, weight, systolic blood pressure, diastolic blood pressure, serum haemoglobin, size of primary tumour, index of tumour stage and histolic grade, serum prostatic acid phosphatase) and four categorical variables with various numbers of levels (performance rating, cardiovascular disease history, electrocardiogram code, bone metastases)
- Some missing data: 62 missing values ($\approx 1\%$)

We forget the classes (Stages of the desease) for performing clustering

Questions

- How many clusters?
- Which partition?

⁵Byar DP, Green SB (1980): Bulletin Cancer, Paris 67:477-488

Data upload without preprocessing

MASSICCC Dashboar	rd Hel	р									1	Profile Lo	gout	
OVERVIEW	Age Cont		Wt Contir •	PF Categ		HX Categ T	SBP	DBP	EKG	HG	\$2 •	Z Contir •	SG Contil	A
FILES	•	_												•
INPUTS	Save													
RESULTS	Preview	<i>i</i>												
		Age	Wt	PF	HX	SBP	DBP	EKG	HG	sz	SG	AP	BM	
	0	75	76	1	1	15	9	5	138	1.4142	8	1.0986	1	
	1	76	?	?	?	?	?	?	?	5.3852	9	2.4849	?	
	2	54	116	1	1	13	7	4	146	6.4807	?	1.9459	1	
	3	69	102	1	2	14	8	5	134	1.7321	9	1.0986	1	
	4	66	?	?	?	?	?	?	?	1.0000	9	2.3979	?	

Run clustering analysis

MASSICCC Dash	aboard Help				Profile	Logout
OVERVIEW	INPUTS					
	Parameters					
	Title	Run demo on c	ancer data set			
RESULTS	Data File	MixtComp-Ex	ample.csv			
	Package	MixMod	MixtComp	BlockCluster		
	Function	Cluster				
	Labels Column			0		
	Cluster Groups	1-7		0		
		Create				
		Create				

It is running on the (Inria) cloud...

MASSICCC Dashboar	d Help				Profile Lo	gout
OVERVIEW	RESULTS					
FILES	Select a job e	ecution fr	om the list below			
INPUTS	03	١	Run Demo On Cancer Data Set MixtComp-Example.csv	44%	5 Feb 16:59	\$
RESULTS	02	۵	MixtComp Cluster Functional-Example.csv		3 Feb 19:15	×
	01	۲	Essai Prostate Vendredi Soir MixtComp-Example.csv		3 Feb 19:03	3

Several quick result overviews... without post-processing

	Output				
WERVIEW					▲ Download Results
us		Variables	12		
NPUTS	Models				
	Model		Criterion	Nb Clusters	Error
	Default		ICL(-12239.6) BIC(-12215.2)	2	No error
	Default		ICL(-12260.4) BIC(-12195.8)	3	No error
	Detault		ICL/-12268.6) BIC(-12208.2)	4	Noerror
	Default		ICL(-12305.1) BIC(-12251.4)	5	No error
	Default		NCL(-12375.0) BNC(-12288.9)	6	No error
	Detault		ICL(-12442.7) BIC(-12354.1)	7	No error
	Default		ICL(-12546.1) BIC(-12546.1)	1	Noerror
	-12.23k -12.3k 80 -12.3k 90 -12.3k 91 -12.4k	1	ICL value / Nb of Clusters		
	5				

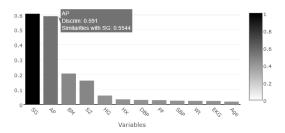
Variable significance on global partition

Importance

This chart represents the discriminating level of each variable. A high value (close to one) means that the variable is highly discriminating. A low value (close to zero) means that the variable is poorly discriminating. Click on one of the bars to display the distribution of this variable and, to also display the similarities between this variable and all the others. The color of the bars reflects the similarities between all the variables and the selected variable. Read more O

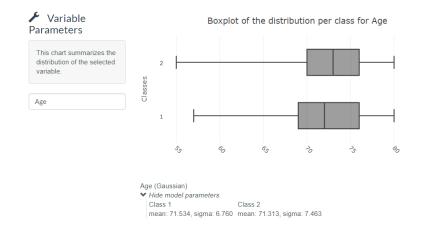
Sort Variables : 1



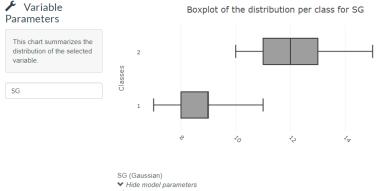


08 0.4 08 24 - 2 1.1.

Variable "Age" difference between clusters

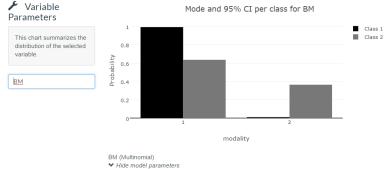


Variable "SG" difference between clusters



Class 1 Class 2 mean: 8.940, sigma: 1.154 mean: 12.087, sigma: 1.405

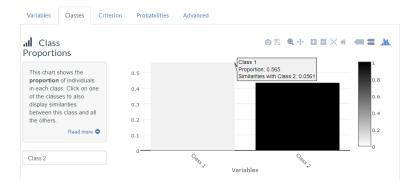
Variable "BM" difference between clusters



Class 1 Class 2

scatter: [0.993,0.007] scatter: [0.633,0.367]

Individual cluster separation (with the cluster weight)



Two strategies in competition

■ Strategy "mice⁶ + MixtComp": MixtComp on the dataset completed by mice

> data.imp=mice(data)

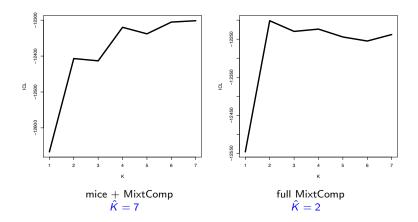
> data.comp.mice=complete(data.imp)

Strategy "full MixtComp": MixtComp on the observed (no completed) dataset

	Partition quality with $K = 2$									
% misclassified 12.8 8.1		Strategy mice + MixtComp full MixtComp								
		% misclassified	12.8	8.1						

⁶http://cran.r-project.org/web/packages/mice/mice.pdf

Choosing K with the ICL criterion



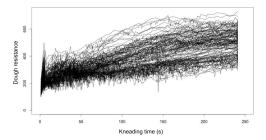
... may lose some cluster information when imputation before clustering

Scoring cancer data following the clustering task

	MASSICCC Dashboard Help	Profile Logout
OVERVIEW	INPUTS	
RLCS	Parameters	
INPUTS	Title	Scoring following the clustering task
RESULTS	Data File	MitComp-Example.csv *
	Package	MixtMod MixtComp BlockCluster
	Function	Predict • Cluster
	Classification Model	Old Windown Carlor DVII Set SFeb 16:59 MatComp-Example.cov
		02 MistCorp Cluster 3 Peb 19:15 Functional-Example.csv
		01 Essai Prostate Verdired Sair 3 Peb 19:03 MixComp-Example.cov
		Greate

Curve "cookies" data set

The Kneading dataset comes from Danone Vitapole Paris Research Center and concerns the quality of cookies and the relationship with the flour kneading process⁷. There are 115 different flours for which the dough resistance is measured during the kneading process for 480 seconds. One obtains 115 kneading curves observed at 241 equispaced instants of time in the interval [0; 480]. The 115 flours produce cookies of different quality: 50 of them have produced cookies of good quality, 25 produced medium quality and 40 low quality.



Upload curves data

	MASSICCC Dashbaard Help	Profile	Logout
overview	Set * Al © ec.0.1.34 As Cotegorical • Apple		
FILES	Upload a file with allist of datatypes for each column.		
INPUTS	Choisisse: un fichier. Aucun fichier choisi		
RESULTS	Function		
	Functional		,
	Sue		
	Preview		
	Function		
	0 025122620216959422257.61097125343,4:263.758		
	1 0.241.129520478231.2.245.716088727869.4.250.18		
	2 0:194.07006418218.2:196.013131806268.4:197.956		
	3 0:137.021447956417,2:154.635389904923.4:170.65		
	4 0.244.120130204111.2.245.627062897663.4.247.13		

Conclusion

Run a clustering task with three clusters

MASSICCC Dashboard Help	Profile Logout
INPUTS	
Parameters	
Title	Clustering of cookies into three clusters
Data File	Functional-Example.csv .
Package	MixMod MixtComp
Function	Cluster •
Labels Column	0
Cluster Groups	3 0
Variable Params	/
	Create

Overview of the three clusters of cookies

MASSICCC Darboard Help Profile Logo	gas. ·
Variable Entropy Class Entropy Parameters Cristrics Ret Variable Similarities Class Similarities	
Cick on one of the boxes to display the model's parameters for the variable.	
Variables	
Cases 3, pr = (0.371)	
Cana 2, pi = (0.27)	
Cause 1, pt = (0.10)	
Hear and 55% CI curves are dates for Function Hear and 55% CI curves are dates for Function Hear and for the function Hear and	

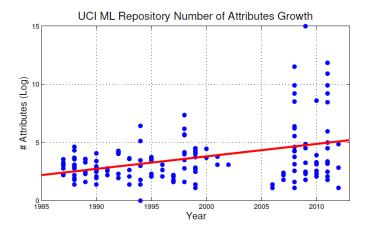




- 2 Model-based clustering
- **3** Mixmod in MASSICCC
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCC



High-dimensional (HD) data⁸



⁸S. Alelyani, J. Tang and H. Liu (2013). Feature Selection for Clustering: A Review. *Data Clustering:* Algorithms and Applications, **29**

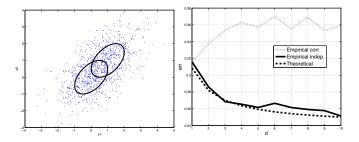
50/79

Bias/variance in HD: reduce variance, accept bias

A two-component *d*-variate Gaussian mixture with intra-dependency:

$$\pi_1 = \pi_2 = rac{1}{2}, \quad {f X}_1 | z_{11} = 1 \sim {\sf N}_d({f 0}, {f \Sigma}), \quad {f X}_1 | z_{12} = 1 \sim {\sf N}_d({f 1}, {f \Sigma})$$

- Each variable provides equal and own separation information
- Theoretical error decreases when d grows: $err_{theo} = \Phi(-\|\mu_2 \mu_1\|_{\Sigma^{-1}}/2)$
- Empirical error rate with the (true) intra-correlated model worse with d
- Empirical error rate with the (false) intra-independent model better with d!



Some alternatives for reducing variance

- Dimension reduction in non-canonical space (PCA-like typically)
- Dimension reduction in the canonical space (variable selection)
- Model parsimony in the initial HD space (constraints on model parameters)

But which kind of parsimony?

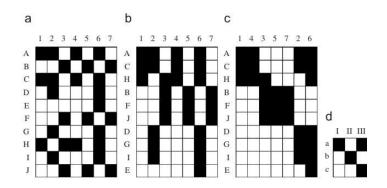
- \blacksquare Remember that clustering is a way for dealing with large n
- Why not reusing this idea for large d?

Co-clustering

It performs parsimony of row clustering through variable clustering

Conclusion

From clustering to co-clustering



[Govaert, 2011]

Notations

- *z_i*: the cluster of the row *i*
- w_j : the cluster of the column j
- **(** z_i , w_j): the block of the element x_{ij} (row *i*, column *j*)
- **z** = (z_1, \ldots, z_n) : partition of individuals in K custers of rows
- $\mathbf{w} = (\mathbf{w}_1, \dots, \mathbf{w}_d)$: partition of variables in L clusters of columns
- **(**z, **w**): bi-partition of the whole data set x
- \blacksquare Both space partitions are respectively denoted by $\mathcal Z$ and $\mathcal W$

Restriction

All variables are of the same kind (research in progress for overcoming that...)

The latent block model (LBM)

- Generalization of some existing non-probabilistic methods
- Extend the latent class principle of local (or conditional) independence
- Thus x_{ij} is assumed to be independent once z_i and w_j are fixed $(\alpha = (\alpha_{kl}))$:

$$p(\mathbf{x}|\mathbf{z}, \mathbf{w}; \boldsymbol{\alpha}) = \prod_{i,j} p(x_{ij}; \boldsymbol{\alpha}_{z_i w_j})$$

• $\pi = (\pi_k)$: vectors of proba. π_k that a row belongs to the *k*th row cluster • $\rho = (\rho_k)$: vectors of proba. ρ_k that a row belongs to the *l*th column cluster • Independence between all z_i and w_j

Extension of the traditional mixture model-based clustering ($\alpha = (\alpha_{kl})$):

$$\mathsf{p}(\mathsf{x}; \boldsymbol{\theta}) = \sum_{(\mathsf{z}, \mathsf{w}) \in \mathcal{Z} \times \mathcal{W}} \prod_{i, j} \pi_{z_i} \rho_{w_j} \mathsf{p}(x_{ij}; \boldsymbol{\alpha}_{z_i w_j})$$

Distribution for different kinds of data

[Govaert and Nadif, 2014] The pdf $p(\cdot; \alpha_{z_iw_i})$ depends on the kind of data x_{ij} :

- Binary data: $x_{ij} \in \{0, 1\}$, $p(\cdot; \alpha_{kl}) = \mathcal{B}(\alpha_{kl})$
- Categorical data with *m* levels: $\mathbf{x}_{ij} = \{x_{ijh}\} \in \{0, 1\}^m$ with $\sum_{h=1}^m x_{ijh} = 1$ and $p(\cdot; \boldsymbol{\alpha}_{kl}) = \mathcal{M}(\boldsymbol{\alpha}_{kl})$ with $\boldsymbol{\alpha}_{kl} = \{\alpha_{kjh}\}$
- Count data: $x_i^j \in \mathbb{N}$, $p(\cdot; \alpha_{kl}) = \mathcal{P}(\mu_k \nu_l \gamma_{kl})^9$
- Continuous data: $x_i^j \in \mathbb{R}$, $p(\cdot; \alpha_{kl}) = \mathcal{N}(\mu_{kl}, \sigma_{kl}^2)$

⁹The Poisson parameter is here split into μ_k and ν_l the effects of the row k and the column l respectively and γ_{kl} the effect of the block kl. Unfortunately, this parameterization is not identifiable. It is therefore not possible to estimate simultaneously μ_k , ν_l and γ_{kl} without imposing further constraints. Constraints $\sum_k \pi_k \gamma_k n_k = \sum_l \rho_l \gamma_{kl} = 1$ and $\sum_k \mu_k = 1$, $\sum_l \nu_l = 1$ are a possibility.

Conclusion

Extreme parsimony ability

Model	Number of parameters
Binary	$dim(\pi) + dim(ho) + KL$
Categorical	$\dim(\pi) + \dim(\rho) + \mathit{KL}(m-1)$
Contingency	$dim(\pi) + dim(ho) + \mathit{KL}$
Continuous	$\dim(\pi) + \dim(\rho) + 2KL$

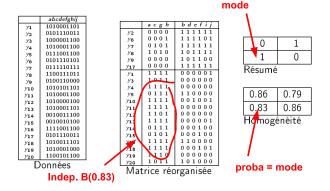
Very parsimonious so well suitable for the (ultra) HD setting nb. param._{HD} = nb. param._{classic}
$$\times \frac{L}{d}$$

Other advantage: stay in the canonical space thus meaningful for the end-user

Conclusion

Binary illustration: easy interpretation

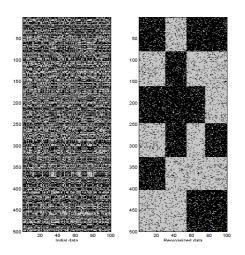
[Govaert, 2011]



Conclusion

Binary illustration: user-friendly visualization

[Govaert, 2011]



$$n = 500, d = 10, K = 6, L = 4$$

MLE estimation: log-likelihood(s)

- Remember Lesson 3: first estimate θ , then deduce estimate of (z, w)
- Observed log-likelihood: $\ell(\theta; \mathbf{x}) = \ln p(\mathbf{x}; \theta)$
- MLE:

$$\hat{oldsymbol{ heta}} = rg\max_{oldsymbol{ heta}} \ell(oldsymbol{ heta};oldsymbol{x})$$

Complete log-likelihood:

$$\ell_c(\boldsymbol{\theta}; \mathbf{x}, \mathbf{z}, \mathbf{w}) = \ln p(\boldsymbol{x}, \boldsymbol{z}, \boldsymbol{w}; \boldsymbol{\theta})$$

=
$$\sum_{i,k} z_{ik} \log \pi_k + \sum_{k,l} w_{jl} \log \rho_l + \sum_{i,j,k,l} z_{ik} w_{jl} \log p(x_i^j; \boldsymbol{\alpha}_{kl})$$

Be careful with asymptotics...

If $\ln(d)/n \to 0$, $\ln(n)/d \to 0$ when $n \to \infty$ and $d \to \infty$, then the MLE is consistent [Brault *et al.*, 2017]

Conclusion

MLE estimation: EM algorithm

E-step of EM (iteration *q*):

$$Q(\theta, \theta^{(q)}) = E[\ell_c(\theta; \mathbf{x}, \mathbf{z}, \mathbf{w}) | \mathbf{x}; \theta^{(q)}]$$

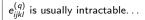
=
$$\sum_{i,k} \underbrace{p(z_i = k | \mathbf{x}; \theta^{(q)})}_{t_{ik}^{(q)}} \ln \pi_k + \sum_{j,l} \underbrace{p(w_i = l | \mathbf{x}; \theta^{(q)})}_{\mathbf{s}_{jl}^{(q)}} \ln \rho_l$$

+
$$\sum_{i,j,k,l} \underbrace{p(z_i = k, w_j = l | \mathbf{x}; \theta^{(q)})}_{e_{ijk}^{(q)}} \ln p(x_{ij}; \alpha_{kl})$$

• M-step of EM (iteration q): classical. For instance, for the Bernoulli case, it gives

$$\pi_k^{(q+1)} = \frac{\sum_i t_{ik}^{(q)}}{n}, \quad \rho_l^{(q+1)} = \frac{\sum_j s_{jl}^{(q)}}{d}, \quad \alpha_{kl}^{(q+1)} = \frac{\sum_{i,j} e_{ijkl}^{(q)} x_{ij}}{\sum_{i,j} e_{ijkl}^{(q)}}$$

MLE: intractable E step



- Consequence of dependency between \mathbf{x}_{ij} s (link between rows and columns)
- Involve KⁿL^d calculus (number of possible blocks)
- Example: if n = d = 20 and K = L = 2 then 10^{12} blocks
- Example (cont'd): 33 years with a computer calculating 100,000 blocks/second

Alternatives to EM

■ Variational EM (numerical approx.): conditional independence assumption

$$p(\mathbf{z}, \mathbf{w} | \mathbf{x}; \boldsymbol{\theta}) \approx p(\mathbf{z} | \mathbf{x}; \boldsymbol{\theta}) p(\mathbf{w} | \mathbf{x}; \boldsymbol{\theta})$$

SEM-Gibbs (stochastic approx.): replace E-step by a S-step approx. by Gibbs

 $\mathbf{z}|\mathbf{x},\mathbf{w};\boldsymbol{\theta}$ and $\mathbf{w}|\mathbf{x},\mathbf{z};\boldsymbol{\theta}$

MLE: variational EM (1/2)

- Use a general variational result from [Hathaway, 1985]
- Maximizing $\ell(\theta; \mathbf{x})$ on θ is equivalent to maximize $\tilde{\ell}_c(\theta; \mathbf{x}, \mathbf{e})$ on (θ, \mathbf{e})

$$\tilde{\ell}_c(\boldsymbol{\theta}; \mathbf{x}, \mathbf{e}) = \sum_{i,k} t_{ik} \ln \pi_k + \sum_{j,l} s_{jl} \ln \rho_l + \sum_{i,j,k,l} e_{ijkl} \ln p(\mathsf{x}_{ij}; \boldsymbol{\alpha}_{kl})$$

where $\mathbf{e} = (e_{ijkl}), e_{ijkl} \in \{0, 1\}, \sum_{k,l} e_{ijkl} = 1, t_{ik} = \sum_{j,l} e_{ijkl}, s_{jl} = \sum_{i,k} e_{ijkl}$

- Of course maximizing $\ell(\boldsymbol{\theta}; \boldsymbol{x})$ or $\tilde{\ell}_c(\boldsymbol{\theta}; \boldsymbol{x}, \mathbf{e})$ are both intractable
- ldea: restriction on **e** to obtain tractability $\mathbf{e}_{ijkl} = t_{ik}s_{jl}$
- New variables are thus now $\mathbf{t} = (t_{ik})$ and $\mathbf{s} = (s_{jl})$
- As a consequence, it is a maximization of a lower bound of the max. likelihood

 $\max_{\boldsymbol{\theta}} \ell(\boldsymbol{\theta}; \boldsymbol{x}) \geq \max_{\boldsymbol{\theta}, \mathbf{t}, \mathbf{s}} \tilde{\ell}_{c}(\boldsymbol{\theta}; \boldsymbol{x}, \mathbf{e})$

MLE: variational EM (2/2)

Approximated E-step

$$Q(\theta, \theta^{(q)}) \approx \sum_{i,k} t_{ik}^{(q)} \ln \pi_k + \sum_{j,l} s_{jl}^{(q)} \ln \rho_l + \sum_{i,j,k,l} t_{ik}^{(q)} s_{jl}^{(q)} \ln p(x_{ij}; \alpha_{kl})$$

- We called it now VEM
- Also known as mean field approximation
- Consistency of the variational estimate [Brault et al., 2017]

Conclusion

MLE: local maxima

- More local maxima than in classical mixture models
- It is a consequence of many more latent variables (blocks)
- Thus: either many VEM runs, or use the SEM-Gibbs algorithm

MLE: SEM-Gibbs

- We have already seen the SEM algorithm in Lesson 3 (thus we do not detail more)
- It limits dependency to starting point, so it limits local maxima
- The S-step: a draw $(\boldsymbol{z}^{(q)}, \boldsymbol{w}^{(q)}) \sim p(\boldsymbol{z}, \boldsymbol{w} | \boldsymbol{x}; \boldsymbol{\theta}^{(q)})$ instead an expectation
- But it is still intractable, thus use a Gibbs algorithm to approx. this draw

Approximated S-step

Two easy draws

$$\mathbf{z}^{(q)} \sim \mathsf{p}(\mathbf{z} | \mathbf{w}^{(q-1)}, \mathbf{x}; \mathbf{ heta}^{(q)})$$

and

$$\boldsymbol{w}^{(q)} \sim p(\boldsymbol{w}|\boldsymbol{z}^{(q)}, \boldsymbol{x}; \boldsymbol{\theta}^{(q)})$$

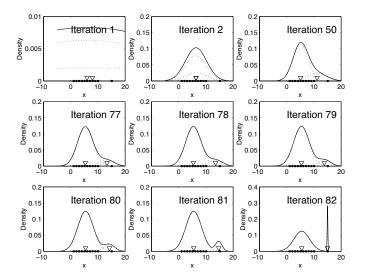
- Rigorously speaking, many draws within the S-step should be performed
- Indeed, Gibbs has to reach a stochastic convergence
- In practice it works well while saving computation time

MLE: degeneracy

- More degenerate situations than in classical mixture models
- It is again a consequence of many more latent variables (blocks)
- The Bayesian regularization (instead MLE) can be an answer

Conclusion

Illustration of a degenerate situation



Conclusion

Bayesian estimation: pitch

 \blacksquare Everything passes by the posterior distribution of $\pmb{\theta}$

$$\mathsf{p}(m{ heta}|m{x}) \propto \underbrace{\mathsf{p}(m{x}|m{ heta})}_{\mathsf{log-likelihood}} \underbrace{\mathsf{p}(m{ heta})}_{\mathsf{prior}}$$

• Then, take (for instance) the MAP as a θ estimate (use a VEM like algo...)

$$\hat{oldsymbol{ heta}} = rg \max_{oldsymbol{ heta}} \mathsf{p}(oldsymbol{ heta} | oldsymbol{x})$$

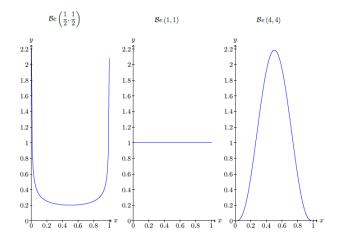
Bayesian estimation: limiting degeneracy

- Interest for avoiding degeneracy is the prior: it acts as a penalization term
- Typical choices are Dirichlet for π and ρ (with independence between π , ρ , α)

$$\mathsf{p}(\boldsymbol{\theta}) = \underbrace{\mathsf{p}(\boldsymbol{\pi})}_{D_{K}(a,...,a)} \times \underbrace{\mathsf{p}(\boldsymbol{\rho})}_{D_{L}(a,...,a)} \times \underbrace{\mathsf{p}(\boldsymbol{\alpha})}_{\mathsf{model dependen}}$$

- The Dirichlet distribution is conjugate, thus easy calculus
- Control degeneracy frequency with the *a* value:
 - **a** = 1: uniform prior, so $\hat{\theta}$ is strictly the MLE (no regularisation)
 - a = 1/2: Jeffreys prior, classical (no informative prior) but may favor degeneracy
 - **a** = 4: a rule of thumb working well for limiting degeneracy frequency

Bayesian estimation: prior overview



Document clustering (1/2)

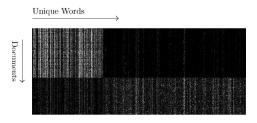
- Mixture of 1033 medical summaries and 1398 aeronautics summaries
- Lines: 2431 documents
- Columns: present words (except stop), thus 9275 unique words
- Data matrix: cross counting document×words

Poisson model

Conclusion

Document clustering (2/2)

 $\mathsf{p}(\hat{\mathsf{z}} \neq \mathsf{z}) \leq 2n \exp \Big\{ -\frac{1}{8} d \Big[\underbrace{\min_{k \neq k'} |\tau_k - \tau_{k'}|}_{\text{overlap}} \Big] \Big\} + \mathcal{K}(1 - \min_{k} \pi_k)^n$



Results with 2×2 blocs

	Medline	Cranfield
Medline	1033	0
Cranfield	0	1398

Conclusion

Running BlockCluster

Configuration

If you change the configuration of your job and save it, it will start a new process with the updated parameters. This will erase previous results.

Parameters		
Title	Trial BlockCluster	
Data File	Blockcluster-Example.csv	
Data Type	Categorical	θ
Rows Cluster Groups	1:5	θ
Column Cluster Groups	1:5	θ
	Update	

Running BlockCluster

MASSICCC Dashboard	Help			Profile I	Logout	
RESULTS	ESULTS					
DATAFILES	Select a job exec	ution from	n the list below			
CREATE JOB	69		Trial BiockCluster Biockcluster-Example.csv 42%	23 May 20	647 🌣	
	68	۲	Genes K1-12 log.cpm.bd	23 May 08	:12 🗸	
	67	1	Genes log.cpm.bd	22 May 15	.38 🗸	•
	65	۲	Genes K1-10 log.cpm.bd	22 May 15	27 🗸	•

Running BlockCluster

Model	Criterion	Nb Clusters	Error	
pik_rhol_multi	ICL (-45557.1)	[2,3]	No error	
pik_rhol_multi	ICL (-45563.3)	[3,3]	No error	
pik_rhol_multi	(CL (~45566.6)	[2,4]	No error	
pik_rhol_multi	ICL (-45573.9)	[4,3]	No error	
pik_rhol_multi	(CL (-45574.6)	[5,3]	No error	
pik_rhol_multi	ICL (-45577.7)	[3,4]	No error	
pik_rhol_multi	(CL (-45578.8)	[2,5]	No error	-

Cluster Plot Criterion Plot

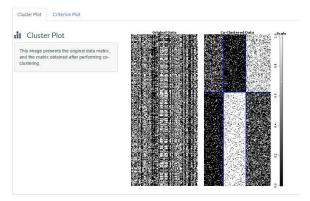
Model Criterion

This chart represents the criterion value for each model that was built. The higher the value (close to 0) the better the model.



Conclusion

Running BlockCluster



Conclusion



1 Introduction

- 2 Model-based clustering
- **3** Mixmod in MASSICCC
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCC





- Use probabilistic modelling as a mathematical guideline
- Use the MASSICCC platform for user-friendly implementation

https://massiccc.lille.inria.fr/