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Validation Française de l'Échelle d'Agressivité et de Colère Compétitive (VFEACC)

French Validation of the Competitive Aggressiveness and Anger Scale (FVCAAS)

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1 ABSTRACT

2 The aim of this study was to propose a French Validation of the Competitive Aggressiveness and Anger Scale (FVCAAS). The instrument was developed from the 3 original version, which is composed of two subscales (six items by subscales) assessing 4 5 aggressiveness and anger in competitive athletes (CAAS, Maxwell & Moores, 2007). Four studies have been conducted with 1,428 competitors. In the first study, the exploratory factor 6 7 analysis extracted the two-factor structure from the original version, both with good internal 8 consistency. The second study confirmed that the two-factor structure of the instrument was 9 consistent with the original version and showed its partial invariance across genders. The 10 third study demonstrated the temporal stability of the FVCAAS. In the fourth study, both concurrent and discriminant validities were confirmed, supporting the validity and reliability 11 of the FVCAAS. The contributions of this study and limitations are discussed, together with 12 13 perspectives for future studies of aggressiveness in competitive sports.

14

15 *Keywords:* Aggressiveness, Validation, Competitive Aggressiveness and Anger Scale,

16 Structural equation modeling, Gender invariance.

17 **RÉSUMÉ**

L'objet de cette étude est de valider une version française de l'échelle d'agressivité et 18 de colère compétitive (VFEACC) parce qu'elle met l'accent sur les antécédents les plus 19 20 importants de l'agression chez l'athlète. L'outil a été développé à partir de la version originale qui est composée de deux sous-échelles de six items mesurant l'agressivité et la colère des 21 compétiteurs (CAAS, Maxwell & Moores, 2007). Quatre études ont été réalisées auprès de 22 1428 compétiteurs. Dans la première étude, l'analyse factorielle exploratoire reproduit le 23 même model à deux factors de la version originale, avec de bonnes consistances internes 24 [Colère: $\alpha = .87$; Agressivité: $\alpha = .84$; Échelle totale: $\alpha = .81$]. La deuxième étude a confirmé 25 la même structure à deux facteurs de la version originale avec de bonnes indices d'ajustement 26 $[X^{2}(53) = 100,77; CFI \text{ et } TLI \ge .95; RMSEA \text{ et } SRMR \le .06].$ Aussi, elle soutient 27 28 l'invariance partielle au niveau le plus contraint (i.e., strict) de la structure factorielle de la VFEACC à travers le genre et montre que cet instrument est valide pour les garçons et les 29 filles. La troisième étude a montré la stabilité temporelle de la VFEACC. Les tests t se sont 30 31 avérés non significatifs, ainsi les scores de corrélation, de consistance et de fiabilité étaient significatifs [$r \ge .60$; $\alpha \ge .69$; ICC $\ge .60$, p < .05]. Enfin, la quatrième étude a montré la 32 validité concourante de l'instrument au travers de ses relations avec d'autres construits 33 mesurant l'agressivité et la colère (Masse, 2001) et le désengagement moral en sport (Corrion 34 et al., 2010), aussi la validité discriminante a été vérifiée par les scores de comparaison selon 35 36 le sexe d'athlète (garçons vs filles) et le type du sport (contact vs sans-contact). Ces résultats supportent la validité, la fiabilité et l'utilité de l'VFEACC pour mesurer l'agressivité et la 37 colère compétitive chez les athlètes français. L'apport de cette étude, certaines limitations et 38 des perspectives pour de futures recherches sur l'agressivité et la colère ont été discutées. 39 *Mots clés* : Agressivité, Validation, Échelle d'Agressivité et de Colère Compétitive, 40 Équation structurelle, Invariance selon le sexe. 41

The current research in applied sport and exercise psychology evidences that the
Competitive Aggressiveness and Anger Scale (CAAS; Maxwell & Moores, 2007) has been
one of the most efficient self-report inventories for measuring athletes' aggressiveness and
anger over the past ten years in studying Chinese and American (Maxell, Moores, & Chowc,
2007; Visek, Maxwell, Watson, & Hurst, 2010), Greek (Bebetsos, Christoforidis, & Mantis,
2008; Bebetsos, 2018), and Portuguese athletes (Sofia & Cruz, 2012, 2017).

The construct validity of the CAAS has been demonstrated among Greek team 48 handball athletes (Bebetsos et al., 2008). The CAAS has also been useful in predicting 49 50 aggressive behaviors or other outcomes in athletes through the relationship between its two subscales and both past unsanctioned aggression and professionalization in Hong Kong 51 52 Rugby players (Maxwell & Vasik, 2009), and athletic experience among Greek male and 53 female water-polo athletes (Bebetsos, 2018). Likewise, Sofia and Cruz (2017) confirmed the relationship between two subscales of the CAAS and anti-social behavior, provocation, 54 aggressive behavior toward opponents and teammates, anger rumination, and state/trait anger 55 56 expression in a sample of Portuguese athletes involved in sports with various contact levels (i.e., roller hockey, rugby, volleyball, basketball, kickboxing, and Greco-Roman wrestling). 57 Little is known, however, about the propensity for aggressiveness and anger among 58 French athletes compared with the other athletes. This dearth of research is partly due to the 59 lack of standard measurement tools that are linguistically, metrically, and conceptually 60 61 equivalent for this sample of athletes. Thus, the research presented here describes the development of a French version of the CAAS (Maxwell & Moores, 2007). 62

A number of sports psychologists have noted that competitive, organized sport is
conducive to unethical and aggressive behavior, particularly if such behavior is positively
reinforced (Conroy, Silva, Newcomer, Walker, & Johnson, 2001) or a win-at-all-costs
philosophy prevails (Keeler, 2007; Nucci & Kim, 2005). Researchers have shown that

67 athletes' aggression is related to a variety of contextual and personal factors, including the type of sport (Maxwell, Visek, & Moores, 2009; Mutz, 2012), training methods (Nosanchuk 68 & MacNeil, 1989), level of competition (Maxwell et al., 2009), sex (Coulomb-Cabagno & 69 70 Rascale, 2005), moral disengagement (Boardley & Kavussanu, 2007; d'Arripe-Longueville, Corrion, Scoffier, Roussel, & Chalabaev, 2010; Stanger, Kavussanu, Boardley, & Ring, 71 2013), performance in contact sports (McCarty & Kelly, 1978; Sheldon & Aimar, 2001), and 72 goal orientation (Rascle, Coulomb, & Pfister, 1998; Dunn & Dunn, 1999). 73 74 Over the past 15 years, theories of aggression have emerged from a wide range of social 75 and psychological perspectives (see Bredemeier, 1985; Kimble, Russo, Bergman, & Galindo, 2010; Maxwell, 2004; Russell, 2008). A number of quantitative studies have used self-report 76 77 measures to examine aggression, particularly to predict aggression in on-field athletes 78 (Kimble et al., 2010). Several instruments are widely used by sport psychologists to study aggression: (a) the Buss-Durkee Hostility Inventory (BDHI; Buss & Durkee, 1957); (b) Buss-79 Perry Aggression Questionnaire (BPAQ, an updated version of the Buss-Durkee Hostility 80 81 Inventory; Buss & Perry, 1992); (c) the Bredemeier Athlete Aggression Inventory (BAAGI; Bredemeier, 1975); (d) the Continuum of Injurious Acts (CIA) and its variants (Bredemeier, 82 1985); and (e) the Sports Behavior Inventory (Conroy et al., 2001). 83 Despite their popularity and the fact they have been widely used in the 84 study of aggression, a number of studies have revealed some psychometric problems that may 85 call into question the reliability and the validity of these instruments (Teubel, Asendorpf, & 86

87 Banse, 2011; Maxwell, 2007; Bryant & Smith, 2001; Hennessy & Wiesenthal, 2001;

88 Stephens, 1998;. Wall & Gruber, 1986). At first, some studies reported that the structural

- 89 models based on Buss and Perry's (1992) original four-factor (29-item) scale failed to
- 90 replicate in three independent American (Bryant & Smith, 2001) and the Chinese samples
- 91 (Maxwell, 2007), and achieved only mediocre goodness-of-fit (i.e., GFI= .76 .81). Some of

92 the scale's items did not reflect the direct endorsement of aggressive traits, had low or 93 multiple loadings, and reverse-scored (Bryant & Smith, 2001). Additionally, in the sport context, the German version of Buss and Perry's (1992) four-factor (27-item) scale failed to 94 95 predict both aggressive behavior in players (i.e., fouls in ten minutes), and other outcomes such as scoring in ten minutes, and coach's judgment on game performance in semi-96 professional basketball (Teubel et al., 2011). The autors indicated that the poor validity of 97 this measure was one of the reasons for their results and confirmed the previous critiques. 98 They noted that, "At the first, standared aggressivness measures could be criticized for their 99 100 poor validity in sport context (Mawwell & Moores, 2007)" (Teubel et al., 2011, p. 397). In addition, the BAAGI (Bredemeier, 1975) is criticized for two main reasons: (1) analysis of 101 102 the internal consistency revealed somewhat poor values for the instrumental subscale (Wall & 103 Gruber, 1986), and (2) its items failed to account for differences across sports aggression (Stephens, 1998; Hennessy & Wiesenthal, 2001). 104 In fact, Maxwell and Moores (2007) made three important criticisms of these 105 106 questionnaires, arguing that they (a) lacked validity, (b) were not sensitive to sport-specific confrontations, and (c) captured moods rather than traits. They indicated that "application of 107 108 the Buss–Durkee and Buss–Perry scales is problematic in sport because some of the items refer to acts that are integral to performance in some sports and other items are not 109 applicable" (Mawwell & Moores, 2007, p. 181). They also noted that aggression in athletes 110 111 has been linked to anger arising, which is linked to increased arousal during the competition and acceptance of aggression; yet none of the existing questionnaires directly measured these 112 factors of arousal arising and acceptance of aggression in sport. To overcome these problems, 113 114 Maxwell and Moores (2007) suggested focusing on the most important antecedents of the aggression: anger and aggressiveness. Aggressiveness is defined as the "acceptance of and 115 willingness to use abuse, both physical and verbal, to gain a competitive advantage'' 116

117	(Maxwell & Moores, 2007, p. 185), and "anger can be triggered by a multitude of events that
118	may occur externally (e.g., threatening language from an opponent or poor officiating) or
119	internally (e.g., negative self-talk related to poor performance and memories of past defeat)"
120	(Maxwell et al., 2009, p. 289).

As a result, Maxwell and Moores (2007) developed the Competitive Anger and 121 Aggression Scale (CAAS) on this basis to measure different antecedents of aggression in 122 sport (i.e., the disposition to become aggressive, acceptance of and willingness to use 123 aggression, incidents of irritation associated with losing, and negative emotions directed at 124 125 opponents). In their study, the CAAS was developed through several phases. It began with constructing a pool of potential items rated on five-point Likert scales ranging from "not at 126 all severe" (1) to "extremely severe" (5). Next, an Exploratory Factor Analysis (EFA) was 127 128 used to evaluate the underlying structure of the items and two factors were extracted: anger 129 and aggressiveness. Six items loaded saliently on each factor, which had a good internal consistency (Anger: $\alpha = .87$; Aggressiveness: $\alpha = .84$; Total: $\alpha = .87$). Subsequently, a 130 Confirmatory Factor Analysis (CFA) revealed that all the parameters had a good fit. The two 131 subscales and the total scale also showed good test-retest reliability. Finally, the concurrent 132 and discriminate validity were also supported, the CAAS score was positively associated with 133 aggression as a personality trait as measured by the BPAQ, and varied across genders and 134 135 sports. They showed that males scored higher than females on both anger and aggressiveness 136 measures; the intensity of anger was relatively low, compared with aggressiveness, but endorsement by British athletes was higher. These analyses confirmed that the CAAS was a 137 suitable instrument for measuring anger and aggressiveness in competitive sports. Hence, the 138 139 CAAS, by its psychometrical quality, was tested and encouraged in cross-cultural studies, which showed that the two-factor structure of the 12-item CAAS was largely replicable 140

across cultures (Maxell *et al.*, 2007; Visek *et al.*, 2010; Bebetsos *et al.*, 2008; Bebetsos, 2018;
Sofia & Cruz, 2012, 2017).

In conclusion, the evidence to date suggests that the CAAS was developed to measure 143 antecedents of aggression in sport (i.e., the disposition to become aggressive, acceptance of 144 145 and willingness to use aggression, incidents of irritation associated with losing, and negative emotions directed at opponents), is a valid measure of anger and aggressiveness in many 146 sports, and that these concepts reliably differentiate players who admit unsanctioned 147 148 aggression from those who do not (e.g., Maxell et al., 2007; Visek et al., 2010; Bebetsos et al., 2008; Bebetsos, 2018; Sofia & Cruz, 2012, 2017). Therefore, the CAAS will be more 149 appropriate in a variety of sports in measuring the traits of aggressiveness and anger than 150 other standared aggressiveness measures that have been criticized for their poor validity in 151 sports (e.g., Maxwell & Moores, 2007; Teubel et al., 2011; Wall & Gruber, 1986; Hennessy 152 153 & Wiesenthal, 2001).

However, according to Maxwell and Moores' critics, it appears that a valid
questionnaire for measuring anger and aggressiveness does not actually exist among French
competitive athletes. The aim of this study, therefore, was to test a French Validation of the
Competitive Aggressiveness and Anger Scale (FVCAAS) based on the original CAAS,
using a large sample of French competitive athletes.

In accordance with Vallerand's (1989) recommendations, the FVCAAS was developed through four studies: (a) development of a preliminary version of the FVCAAS (Study 1); (b) confirmation of the factor structure of the FVCAAS and testing its gender invariance (Study 2); (c) evaluation of the temporal stability of the instrument (Study 3); and (d) assessment of its concurrent and discriminant validity (Study 4).

164 1. Study 1

165 The aims of the first study were (a) developing items for a preliminary FVCAAS, (b) 166 verifying the clarity of items using a sample of French athletes, and (c) assessing its structure 167 using EFA.

168 *1.1. Method*

169 *1.1.1. Participants and procedure*

170 Sample 1a consisted of 30 athletes (15 men; 15 women; $M_{age} = 15.66$ years; SD =1.47; range: 14-18) who were randomly selected from various individual and team sports; 171 172 the sample was used to assess the clarity of the items of the FVCAAS. Sample 1b consisted of 201 volunteer athletes (109 males; 92 females; $M_{age} = 19.61$; SD = 4.95; range: 173 14-25) in a range of sports (e.g., handball, football, basketball, and judo). These athletes 174 provided data for the EFA. They completed the initial FVCAAS at their clubs. Athletes 175 under 18 were enrolled if they agreed to participate and their parents provided written consent 176 177 for their participation, in line with ethical standards. Participation was voluntary and anonymous and participants were assured that their data would remain strictly confidential. 178

179 *1.2. Results and Discussion*

180 *1.2.1. Formulation of items*

First, in accordance with Vallerand's (1989) and Brislin's (1986) back-translation procedures, the CAAS was translated into French. Two independent translations into French, led by two English-language experts, were undertaken to establish the best expression of each item, and a French version was created thereafter. The French version was then translated back into English by a perfectly bilingual speaker. The French version was established based on this process of translating and adapting the instruments. The original CAAS and the French version FVCAAS were equivalent (Table 1). Responses to the FVCAAS items were

made on a five-point Likert scale (1 = almost never; 2 = occasionally; 3 = sometimes; 4 =
quite often; 5 = almost always).

190 *1.2.2. Assessment of item clarity*

In the second phase, the clarity of the 12-item FVCAAS was measured using a sixpoint Likert scale ranging from 1 ("not at all clear") to 6 ("completely clear"). Sample 1a was asked to respond as honestly as possible to the questions and assured that their responses would be anonymous. The results showed that the items had satisfactory clarity (M = 4.88, SD = 1.12) and in one-to-one interviews all the participants declared that the items were easy to understand and represented their behaviors.

197 *1.2.3. Exploratory factor analysis*

In the third phase, responses to the FVCAAS were analyzed using SPSS12.0 software.
Principal component analysis followed by Oblimin rotation with Kaiser normalization
revealed two six-item factors that together accounted for 45.90% of the variance in scores.
The two factors were named "Anger" and "Aggressiveness," and their initial Eigen values
were 4.14 and 1.36, respectively; they accounted for 34.45 % and 11.38% of variance in
scores, respectively. Also, factor loadings generally were good to high (Table 1).

Table 1 also shows that in descriptive terms the mean Anger score (M = 2.64, SD =204 1.23) was higher than the mean Aggressiveness score (M = 1.94, SD = 1.23). Our results are 205 in accord with Maxwell et al.'s (2009) study on Hong Kong Chinese athletes, which showed 206 207 that the intensity of anger was relatively high, compared with the aggressiveness in different contact sports (i.e., basketball, rugby union, association football/soccer, and squash). Our 208 results, however, did not support the early findings in British athletes (Maxwell & Moores, 209 210 2007). Our findings suggest that the expression of Anger may be more frequent than Aggressiveness among French athletes. Our results also support the idea that intense anger 211 and irritation, which are generated by losing frequently or a referee's mistakes, may be 212

reflected as an ineffective reaction directed at opponents or officials in French contact sports, 213 such as soccer (Maxwell et al., 2009; Reynes, Canovas, Ferrand, & Pantaleon, 2008). 214 All items had acceptably low skewness and kurtosis ($< \pm 1.98$), and the distribution of 215 data did not deviate significantly from normality. Both subscales and the scale as a whole 216 showed good internal consistency, as measured by Cronbach's alpha (Anger: $\alpha = .71$; 217 Aggressiveness: $\alpha = .76$; Total: $\alpha = .82$). 218 219 "(Insert Table I here)" 2. Study 2 220 221 The second study aimed to assess (a) the factor structure of the FVCAAS using CFA and (b) the gender invariance of the instrument according to previously described procedures 222 (Chen, 2007; Cheung & Rensvold, 2002; Keith, 2015). 223 2.1. Method 224 2.1.1. Participants and procedure 225 Two new samples of volunteer athletes who were competing regularly were recruited 226 from various clubs. Sample 2a consisted of 252 athletes (141 male; 111 females; $M_{age} =$ 227 18.84; SD = 3.99) and provided data for the first CFA. Sample 2b consisted of 619 athletes 228 (310 males; 309 females; $M_{age} = 16.97$; SD = 2.96) and provided data for the second CFA 229 230 and the invariance test. The athletes completed the FVCAAS using a five-point scale, ranging from 1 ("almost never") to 5 ("almost always"). 231 2.1.2. Data analysis 232 CFAs were conducted using IBM SPSS AMOS 24 software (Arbuckle, 2016) to 233 confirm the factor structure of our scale in a new sample (Sample 2a). The appropriateness of 234 the model must assess with global goodness-of-fit indices to the data. The fit of the model 235 was assessed using five indices, with the criteria for acceptable fit (Hu & Bentler, 1999) 236

237	given in the following parentheticals: chi-squared (χ^2); root mean square error of
238	approximation (RMSEA; \leq .08 at 90% CI); comparative fit index (CFI; \geq .90); the Tucker-
239	Lewis index (TLI; \geq .90); and standardized root mean square residual (SRMR; \leq .08).
240	The gender invariance of the FVCAAS was computed using procedures described
241	previously (Chen, 2007; Cheung & Rensvold, 2002; Keith, 2015). Two perspectives were
242	used: (a) the traditional perspective based on change on chi-squared ($\Delta\chi^2$), and (b) the
243	practical perspective based on differences in the comparative fix index (Δ CFI). The $\Delta\chi^2$ test
244	allows one to compare nested models (e.g., configural vs. metric; metric vs. strong; and
245	strong vs. strict), but due to its sensitivity to model complexity and large samples, it is nearly
246	always large and statistically significant (Chen, 2007). Cheung and Rensvold (2002)
247	recommended ΔCFI over $\Delta \chi^2$ based on its relative insensitivity to model complexity, sample
248	size, and overall fit. They stated, "A value of ΔCFI smaller than or equal to01 indicates that
249	the null hypothesis of invariance should not be rejected" (Cheung & Rensvold, 2002, p. 251).
250	In this study, in deciding whether the fit of more restrictive models was gender invariant we
251	used Chen's (2007) criteria for invariance between nested models given a sample of >300:
252	$\Delta CFI \le .01$; $\Delta RMSEA \le .015$ and $\Delta SRMR \le .03$ for metric invariance; and $\Delta CFI \le .01$,
253	$\Delta RMSEA \leq .015$ and $\Delta SRMR \leq .01$ for the more complex models.

- 254 2.2. Results and Discussion
- 255 2.2.1. Confirmatory factor analysis

To verify the two-factor structure suggested by EFA, we carried out CFA with a new sample. Figure 1 shows that the factor loadings were fair to high (.58-.78 for Sample 2a) and the CFA results provided good fit indices [CFI and TLI \geq .93, RMSEA \leq .08, and SRMR \leq .05] and confirmed that the 12-item FVCAAS has a two-factor structure (Table 2). These

- 260 results confirmed that the FVCAAS was similar to the original CAAS scale (Maxwell &
- 261 Moores, 2007).

"(Insert Figure 1 here)"

263 2.2.2. Gender invariance

Table 2 reveals that the model fits each data group well, suggesting that the following 264 invariance verification was meaningful. Also, it reveals that the configural invariance model 265 was accepted, based on the global fit indices. There was a significant χ^2 value and good fit 266 indices [CFI and TLI \ge .92, RMSEA = .05, and SRMR= .06], which makes it possible to 267 envisage the metric invariance. Next, the metric invariance model was practically supported 268 by the data. The addition of test loading factors reduced the fit according to $\Delta \chi 2$ [i.e., $\Delta \chi 2$ 269 270 =37,58, $\Delta p < .01$], due to its sensitivity to complex models and large samples, but not according to Δ CFI, Δ RMSEA, and Δ SRMR. The change value for these measures was $\leq .01$, 271 which is equal to or well below the recommended criteria (Chen, 2007; Cheung & Rensvold, 272 273 2002), which revealed that the FVCAAS two-factor structure between men and women should be similar. However, we rejected full, strong invariance. The differences in the χ^2 and 274 CFI between the strong and the metric invariance models were very large and beyond the 275 276 recommended criteria [i.e., $\Delta \gamma 2 = 79.97$, $\Delta p < .01$, $\Delta CFI=.03$]. Next, we tested for partial strong invariance. The modification indices put forward by the AMOS program suggested 277 unconstraining the error of measurement for item agg1 (Violent behavior directed towards an 278 opponent is acceptable) and item agg2 (It is acceptable to use illegal physical force to gain an 279 advantage) in both groups, which the analysis showed satisfactory partial strong invariance 280 281 according to changes in CFI and other fit indices [Δ CFI=.01, Δ RMSEA, and Δ SRMR = 0]. Finally, partial strict invariance was also supported by data [Δ CFI=.01, Δ RMSEA= 0, and 282 Δ SRMR = -.01], implying that the test intercepts and test residuals were the same in both 283 284 male and female groups (Chen, 2007; Cheung & Rensvold, 2002), with the exception of two items (agg1 and agg2). Therefore, an adaptation to improve the invariance in these items 285 according to gender should be considered in future studies. In sum, this series of sample 286

analyses indicated partial invariance at the most complex (i.e., strict) levels of the FVCAAS
factorial structure across genders.

289

"(Insert Table II here)"

290 **3. Study 3**

291 The aim of this study was to check the temporal stability and internal consistency of292 the FVCAAS.

293 *3.1. Method*

3.1.1. Participants and procedure

The new sample (Sample 3) consisted of 61 volunteer athletes (42 males, 19 females;

296 $M_{age} = 18.47$; SD = 4.91) who undertook regular physical activity. They completed the

297 FVCAAS twice (T1 and T2) in a four-week period as interval between two tests.

298 *3.2. Results and Discussion*

299 The intraclass correlation coefficient (ICC) and 95% CI for the ICC, together with a series of paired-sample *t*-tests, were used to assess the temporal stability of the 300 questionnaire's structure. The ICCs and the ICC 95% CIs for both subscales are provided in 301 302 Table 3. The ICCs were statistically significant in all cases. The ICCs for Anger, Aggressiveness, and the FVCAAS as a whole were .71, .71, and .79, respectively at T1. The 303 corresponding values at T2 were .72, .60, and .66. These results indicate that both subscales 304 and the FVCAAS had acceptable temporal stability. The results of *t*-tests comparing subscale 305 scores at T1 and T2 were non-significant, indicating that FVCAAS responses remained 306 307 similar over a four-week period. In addition, significant positive correlations were found between the first and second collections of data for each subscale (Anger: r = .63; 308 Aggressiveness: r = .60; Total: r = .62; all ps < .05). These findings show that the FVCAAS 309 and its subscales had good temporal stability over four weeks. Finally, the internal 310

311 consistency of the subscales was assessed at T1 and T2; once again the Cronbach's α were 312 acceptable (T1: $\alpha = .72 - .79$; T2: $\alpha = .69 - .71$).

313

"(Insert Table III here)"

314 **4. Study 4**

Study 4 had two aims. First, to assess the concurrent validity of the FVCAAS by
calculating correlations between FVCAAS scores and scores on two other instruments
measuring constructs that are theoretically related to those measured by the CAAS (Maxwell
& Moores, 2007). The second aim was to check its discriminant validity.

319 In accordance with Maxwell and Moores's (2007) research, the concurrent validity was assessed by comparing subscale and total scores on the FVCAAS with (a) subscale and 320 total scores on the French version of BPAQ (Masse, 2001), which measures four components 321 322 of trait aggression (physical aggression, verbal aggression, anger, and hostility) and (b) an index of moral disengagement, which has been defined as "the self-regulatory process by 323 which individuals cognitively restructure their inhumane conduct, the negative effects of their 324 actions, their role in causing harm, or the targets of their transgressive acts" (d'Arripe-325 Longueville et al., 2010, p.598). Researchers (e.g., Boardley & Kavussanu, 2007; d'Arripe-326 Longueville et al., 2010; Stanger et al., 2013) have demonstrated: (a) strong positive 327 associations between sport-related moral disengagement and both negative affect (e.g., 328 329 irritability) and antisocial or transgressive behavior (i.e., behavior intended to harm or 330 disadvantage other athletes; the intention to cheat) and (b) moderate negative correlations between sport-related moral disengagement and prosocial behavior (i.e., behavior intended to 331 help or benefit other athletes). 332

333 Next, to check the discriminant validity of the FVCAAS, we used a multivariate 334 analysis of variance (MANOVA) to examine how FVCAAS subscale scores varied across 335 genders and sports. Maxwell and Moores (2007) suggested that scores on the Competitive

336	Aggressiveness and Anger subscales of the CAAS varied across genders and sports. They
337	reported that male athletes tended to be more aggressive in sports than female athletes, and
338	that aggressive athletes were drawn to contact sports over non-contact sports.

339 *4.1. Method*

340 *4.1.1. Participants and procedure*

A new sample (Sample 4) of 175 volunteer athletes (106 males; 69 females; M_{age} = 19.59; SD = 5.34; range: 12-35) who were involved in high-contact (rugby: n = 83) or lowcontact (basketball: n = 92) sports (see Table 5). This sample was recruited to verify the concurrent and discriminant validity of FVCAAS. The athletes completed four questionnaires to enable us to assess the correlations between the FVCAAS and (a) a French version of the BPAQ and (b) an index of moral disengagement. All participants were treated with APA ethical guidelines.

348 *4.1.2. Instruments*

349 *4.1.2.1. French Validation of the Competitive Anger and Aggressiveness Scale*

350 The 12-item FVCAAS developed for this research was used. In Sample 4 the 12-item

351 model was a good fit for the data, $\chi^2(53) = 118.48$; p < .01; CFI = .90; TLI = .90; RMSEA =

352 .07; 95% CI RMSEA: .06-.10; SRMR = .06. Values of Cronbach's alpha were good (Anger: α

- 353 = .78; Aggressiveness: α = .81; Total: α = .86).
- 354 4.1.2.2. Buss–Perry's Aggression Questionnaire

The validated French version of the BPAQ (Masse, 2001) has been used to assess the various aspects of aggression in athletes. It consists of 20 items distributed unequally across four subscales; namely, physical aggression, verbal aggression, anger, and hostility. The physical aggression subscale measures one's tendency to be physically aggressive toward other people or objects (sample item: "If I have to resort to violence to protect my rights, I will"). The verbal aggression subscale measures verbal expression of aggression (sample item: "When people annoy me I tell them what I think of them"). The anger subscale
measures the respondent's feelings of anger (sample item: "I sometimes feel like gunpowder
ready to explode"). The hostility subscale assesses hatred of others and of life (sample item:
"When people are especially nice to me, I wonder what they want"). Responses are given
using a five-point Likert scale ranging from "Extremely uncharacteristic of me" (1) to
"Extremely characteristic of me" (5). In this study the subscales had good internal
consistency (α: .75-.90; see Table 4).

368 *4.1.2.3. French Short Moral Disengagement Scale*

369 Moral disengagement was assessed with the French Short Moral Disengagement Scale (Corrion, Scoffier, Gernigon, Cury, & d'Arripe-Longueville, 2010). The scale consists 370 of three items measuring minimization of transgressions and their consequences (e.g., "It's 371 372 not serious if I behave badly [cheating or aggression] if it's in order to win") and three items measuring projection of fault onto others (e.g., "It's not my fault if I behave badly [cheating 373 or aggression] if my opponent started it"). Responses are given on a six-point Likert scale 374 ranging from 1 ("Do not agree at all") to 6 ("Completely agree"). The reliability of the scale 375 as a whole is $\alpha = .88$. 376

377 4. 2. Results and Discussion

378 *4. 2.1. Concurrent validity of the FVCAAS*

379 Correlations were used to evaluate the relationships between subscale and total scores
380 on the FVCAAS and (a) subscale (physical aggression, verbal aggression, anger, and

- hostility) and total scores on the French BPAQ, and (b) moral disengagement. Pearson
- 382 correlations indicated FVCAAS-aggressiveness, FVCAAS-anger and FVCAAS-total were
- 383 systematically correlated with BPAQ-physical aggression, BPAQ-verbal aggression, BPAQ-
- anger, BPAQ-hostility, and BPAQ-total (r: .39-.57, all ps < .05; see Table 4). These findings
- are consistent with the literature (e.g., Maxwell & Moores, 2007).

386	The concurrent validity of the FVCAAS was also demonstrated by the existence of
387	the expected correlations between FVCAAS-aggressiveness, FVCAAS-anger, and FVCAAS-
388	total and moral disengagement (r : .4255, all $ps < .05$; see Table 4). These findings are in line
389	with the other research demonstrating strong positive associations between moral
390	disengagement and both antisocial and aggressive behaviors in sport (e.g., Boardley &
391	Kavussanu, 2007; d'Arripe-Longueville et al., 2010). As the FVCAAS was shown to be
392	related to both the established constructs with which it was compared, we concluded that the
393	instrument had concurrent validity.
394	"(Insert Table IV here)"
395	4. 2.2. Discriminant validity of the FVCAAS
396	To determine how FVCAAS scores varied across genders and sports, we carried out a
397	2x2 between-subject MANOVA, with gender and types of sport as the independent variables
398	and FVCAAS-aggressiveness, FVCAAS-anger, and FVCAAS-total as the dependent
399	variables. Gender affected the results [Wilks's $\lambda = .88$, F (2, 170) = 11.22, $p < .001$, $\eta^2 = .12$]
400	as did the type of sport [Wilks's $\lambda = .85$, F (2, 170) = 5.33, $p < .01$, $\eta^2 = .06$], but there was no
401	interaction between gender and type of sport [Wilks's $\lambda = .96$, F (2, 170) = 0.73, $p = .69$, $\eta^2 =$
402	.004]. Separate univariate ANOVAs were conducted to analyze group differences in each
403	dependent variable. There were gender differences in FVCAAS-aggressiveness [F (1, 171)
404	= 22.56, $p < .001$, $\eta^2 = .16$], FVCAAS-anger [F (1, 171) = 6.82, $p < .01$, $\eta^2 = .04$] and
405	FVCAAS-total [F (1, 171) = 16.75, $p < .001$, $\eta^2 = .11$]. The type of sport affected FVCAAS-
406	aggressiveness [F (1, 171) = 6.71, $p < .05$, $\eta^2 = .09$] but not FVCAAS-total [F (1, 171)
407	= 1.65, $p = .2519$, $\eta^2 = .01$] and FVCAAS-anger [F (1, 171) = 0.03, $p = .86$, $\eta^2 = .01$]. No
408	multivariate outliers were identified and the means are provided in Table 5.
409	As a result, discriminant validity was established based on the athletes' differences. A
410	number of authors have reported that there is a gender difference in the tendency to value

aggressiveness and anger in sport (e.g., Maxwell & Moores, 2007; Christoforidis, Kalivas, 411 Matsouka, Bebetsos, & Kambas, 2010; Bebetsos, 2018). This general observation was 412 completely supported by our data. Male athletes had higher FVCAAS-total, FVCAAS-anger, 413 414 and FVCAAS-aggressiveness scores than female athletes (Maxwell & Moores, 2007). In the present study, male athletes appear to take advantage of their anger more than female 415 athletes. These findings may suggest that male athletes express their anger, though they 416 417 experience persistent anger and therefore had a greater tendency toward expressing both their anger and aggressiveness compared with female athletes. However, our results are not 418 419 consistent with the results of previous research showing no statistically significant differences between female and male athletes regarding trait anger and anger expression styles in 420 421 taekwondo (Lapa, Aksoy, Certel, Çalışkan, Özçelik, & Çelik, 2013) and volleyball (Esfahani, 422 Gheze, & Soflu, 2010). The results of this previous research were somewhat unsurprising given the nature of the psychosocial factors underlying anger, such as frequently losing or a 423 referee's mistakes (Reynes et al., 2008). This also makes sense when taking into account the 424 importance of anger expression for the athlete's performance (e.g., Lapa et al., 2013; 425 Maxwell, 2004; Robazza & Bortoli, 2007), which may play a role in explaining these 426 427 findings. But these are factors beyond the scope of the current study. Further investigation on possible associations between these factors may be warranted. 428 429 As expected, there were also differences in FVCAAS-aggressiveness based on type of 430 sport. Athletes involved in high-contract sports had higher FVCAAS-aggressiveness scores than those involved in low-contact sports. This finding is consistent with previous results 431 (Maxwell & Moores, 2007; Maxwell et al., 2009). It may be that contact sports give players 432

433 more opportunity for legitimate expression of aggression than non-contact sports.

434

"(Insert Table V here)"

435 **5. General Discussion**

The aim of this research was to develop and test a French Validation of the CAAS. 436 First, the 12 items of the CAAS were translated by two English-language experts into French 437 to create a preliminary list of items for the FVCAAS. The EFA indicated that these 12 items 438 439 were organized equally into two factors, Anger and Aggressiveness, with satisfactory internal consistency (Study 1). Study 2 provided evidence of the factorial validity of the FVCAAS, its 440 similarity to the original version, and demonstrated its partial invariance across gender at the 441 most complex level (i.e., strict), also an adaptation to improve the invariance according to 442 gendre was suggsted for two items (agg1 and agg2) in future study. The temporal stability of 443 444 the instrument was also confirmed (Study 3). The concurrent validity of the FVCAAS was demonstrated by relationships between its scales and two instruments measuring constructs 445 that are theoretically related to athletes' aggression. Additionally, the discriminant validity of 446 447 the FVCAAS was demonstrated by the differences in scores associated with genders and sports (Study 4). Together, these analyses provide comprehensive evidence of the validity of 448 the FVCAAS and its ecological validity because the item content was provided by junior and 449 450 adult competitive athletes.

This series of studies carried out with French athletes resulted in the creation of six-451 item scales measuring competitive anger (sample item: "I become irritable if I am 452 disadvantaged during a match") and competitive aggressiveness (sample item: "Opponents 453 accept a certain degree of abuse"), which are structurally and theoretically valid, reliable, 454 455 stable over time, and well-suited for testing assumptions about gender differences in anger and aggressiveness. The links between the FVCAAS and different facets of trait aggression 456 and moral disengagement are in line with the existing sport psychology literature, which 457 458 indicates that competitive anger and aggressiveness are associated with physical and verbal aggression, anger, and hostility (Maxwell & Moores, 2007). It also demonstrates that moral 459 460 disengagement reduces negative emotional reactions to transgressive behavior and increases

461 one's self-reported likelihood of behaving antisocially in sport (e.g., d'Arripe-Longueville *et al.*, 2010; Stanger *et al.*, 2013).

Although the FVCAAS demonstrated good psychometric properties in this research, 463 464 several limitations must be acknowledged. First, all research relying on self-report measures, particularly research concerning moral variables, may be subject to social desirability bias. 465 A second limitation is that the assessment of the scale's construct validity was limited 466 to analysis of its convergent validity and ability to reproduce the gender and type-of-sport-467 related differences that exist in theoretically related constructs. There is, therefore, a need for 468 469 further research into the scale's discriminant validity based on comparisons with scales measuring other constructs. The discriminant validity of the instrument would be 470 471 strengthened by evidence that FVCAAS-anger and FVCAAS-aggressiveness are negatively 472 correlated with theoretically related constructs, such as self-regulatory efficacy in sport (Boardley & Kavussanu, 2007; d'Arripe-Longueville et al., 2010). 473

A third limitation is the specific focus of the FVCAAS. Despite the validity of the 474 FVCAAS for measuring aggressiveness and anger in a variety of sports, some items (e.g., 475 agg1, agg2, and ang6) that described the acceptability or normativity of aggression and anger 476 expression against the opponent and the referee to gain a competitive advantage are more 477 common and suitable in contact sports (e.g., football/soccer, basketball, rugby, and boxing) 478 than non-contact sports (i.e, climbing, swimining, and gyminstics), which should be 479 480 acknowledged (e.g., Conroy et al., 2001; Maxwell & Moores, 2007). Therefore, it would be of interest in future studies to undertake more careful scientific investigations to improve the 481 appropriateness of the aggressiveness subscale according to non-contact sport types, or 482 483 change the format of the original five-point scale from (almost never = 1; occasionally = 2; sometimes = 3; quite often = 4; almost always = 5) to a six-point scale ranging from 1484 ("extremely uncharacteristic of me") to 6 ("extremely characteristic of me") to eliminate the 485

486 scale's midpoint and force respondents to decide whether each statement was characteristic of487 them (Bryant & Smith, 2001).

A fourth limitation is that, to date, no predictive validity data are available for this 488 489 study. The current results of different psychometric tests and their interpretations provide evidence that the FVCAAS is a valid measure of anger and aggressiveness in many sports, 490 and thus there is no coherent argument against its future use with the French athletic 491 population until such data are available to evaluate the test's predictive validity. Therefore, 492 we proposed to continue testing the predictive validity of FVCAAS according to Maxwell 493 494 and Moores', (2007) and Teubel et al.'s (2011) recommendations in future studies, as a second step of validating this scale. The recommendations of Maxwell & Moores (2007) and 495 496 Teubel et al. (2011) indicates the importance of combining both implicit and explicit 497 aggressiveness' questionnaires, interviews, and observations to gain better insight into aggression in sport, rather than relying solely on the subjective interpretation of the athletes. 498 Accordingly, and consistent with prior studies providing the predictive validity of the original 499 500 CAAS scale (e.g., Bebetsos et al., 2008; Maxwell & Vasik, 2009; Sofia & Cruz, 2017; Bebetsos, 2018), we hypothesized the existence of a significant positive correlation between 501 FVCAAS scores and frequency of observed aggressive behavior for players in the premier 502 basketball league in France. It will be important to deepen this analysis in future studies. 503 504 It would also be of interest to compare the levels of anger and aggressiveness between 505 French and British athletes, and to explore whether the cultural differences between France

and Great Britain (Hébert & Dugas, 2010) influence their experiences and expressions of
anger, as well as the use of aggressive tactics in their behaviors to achieve athletic success.
Indeed, many empirical studies have reported that culture was a potential source of difference
in aggressiveness between athletes. For example, Hong Kong Chinese athletes reported lower
aggressive responses than both British (Maxwell *et al.*, 2007) and American athletes (Visek

511 *et al.*, 2010), although the difference was not observed in anger (Maxwell *et al.*, 2009;

512 Maxwell *et al.*, 2007). The anger reflected frustration provoked by losing points or games and
513 an official's mistakes (Reynes *et al.*, 2008).

From a practical standpoint, the FVCAAS could help identify athletes with low or 514 high anger or aggressiveness and determine how these constructs are related to performance 515 among French athletes. Some studies have demonstrated that high aggressiveness and an 516 intermediate level of anger enhance athletic performance, particularly in contact sports (e.g., 517 McCarty & Kelly, 1978; Robazza, & Bortoli, 2007; Sheldon & Aimar, 2001). Anger may 518 519 have a positive or negative effect on performance through its effect on arousal; moderate arousal enhances performance but high arousal has a detrimental effect. These findings 520 suggest that aggressiveness and anger are significant aspects of an athlete's profile and 521 522 should be taken into account to enhance training methods. In conclusion, despite the aforementioned limitations, this series of studies fills a gap 523 in the existing literature and demonstrates that the FVCAAS is a sound, sport-specific 524 instrument for assessing competitive anger and aggressiveness in French athletes. 525

526 **Conflict of interest "none"**

527 The authors declare that they have no links of interest.

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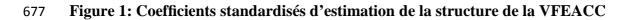
Table I: Les items de la VFEACC et l'échelle originale (CASS, Maxwell & Moores, 2007) 672

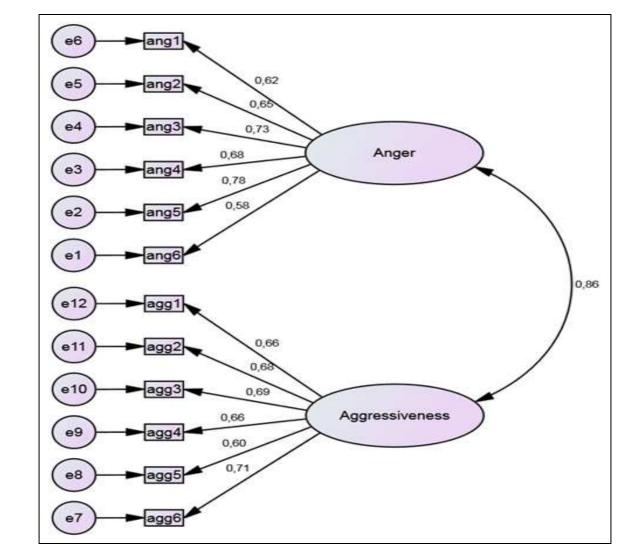
Table I: The items of the FVCAAS and original scale (CASS, Maxwell & Moores, 2007) 673

Items assigned to each factor	Μ	SD	skewness	kurtosis	α	EFA
Anger (F1)	2.64	1.23	-0.70	-0.73	.71	
Ang1. I become irritable if I am disadvantaged during a match.	2.83	1.22	-0.42	-0.97		.58
Je deviens irritable si je me sens dépassé(e) lors de ma prestation.						
Ang2. I feel bitter towards my opponent if I lose.	2.37	1.25	0.64	-0.52		.64
J'en veux à mes adversaires quand je perds.						
Ang3. I get mad when I lose points.	2.27	1.26	0.74	-0. 53		.69
Je deviens fou/folle quand je perds des points.						
Ang4. I show my irritation when frustrated during a game.						
Je montre mon agacement quand je suis frustré(e) pendant ma	2.36	1.26	0.59	-0.77		.55
prestation.						
Ang5. I find it difficult to control my temper during a match.	2.26	1.29	0.73	-0.51		.65
Je trouve difficile de contrôler ma colère pendant ma prestation.						
Ang6. Officials' mistakes make me angry.	3.08	1.07	-0.13	-1.09		.66
Les erreurs des arbitres me font enrager.						
Aggressiveness (F2)	1.94	1.23	0.81	0.72	.76	
Agg1. Violent behavior, directed towards an opponent, is						
acceptable.	1.69	1.07	0.17	0.17		.79
Un comportement violent, dirigé contre un adversaire, est						
acceptable.						
Agg2. It is acceptable to use illegal physical force to gain an						
advantage.	1.96	1.23	1.28	0.38		.78
Il est acceptable d'utiliser une force physique non autorisée pour						
obtenir un avantage.						
Agg3. I taunt my opponents to make them lose concentration.						
Je provoque mes adversaires pour leur faire perdre leur	1.89	1.29	1.29	0.16		.63
concentration.						
Agg4. I use excessive force to gain an advantage.	2.61	1.33	0.26	-1.19		.60
J'utilise une force physique excessive pour obtenir un avantage.						
Agg5. I verbally insult opponents to distract them.	1.35	0.87	1.74	1.86		.71
J'insulte verbalement mes adversaires pour les distraire.						
Agg6. Opponents accept a certain degree of abuse	2.16	1.13	0. 17	-0.57		.55
Les adversaires tolèrent un certain degré de violence.						

674 Notes. α: Cronbach's alpha; EFA: loading in the exploratory factor analysis. Ang: Anger; Agg: Aggressiveness.

675 The items French translations are in italics. For each item the participant had to answer on a five-point Likerttype scale from "almost never" (1) to "almost always" (5), n = 201. 676





678 Figure 1: Standardized coefficients of estimation of the structure of the FVCAAS

Notes. The standardized coefficients of estimation are all significant at p < .001, n = 252.

Table II: Indices d'ajustement des modélisations structurales de la VFEACC.

Table II: Fit indices for the structural modelings of the FVCAAS.

Models	×2	df	d	CFI TLI		RMSEA	90% CI	SRMR	SRMR ECVI	AIC	Comparaison	⊳χ²	∆df	∆p ∆C	ACFI ARMSEA		ASRMR
Structure $(n = 252)$																	
All (12 items)	100.77	53	< .01	96.	.95	90.	.0407	.05	69.	174.76	I	I	I	I	1		I
Anger (6 items)	17.26	6	.05	98.	76.	90.	.0010	.03	.21	53.25	I	I	I	I	I		I
Aggressiveness (6 items)	28.01	6	< .01	96.	.93	.08	.0513	.04	.25	64.01	I	I	I	I	1		Ι
Gender invariance $(n = 619)$																	
Man $(n = 310)$	190.08	53	< .01	.93	.91	.07	.07–.08	.05	.68	211.73	I	I	I	I	1		I
Woman $(n=309)$	137.73	53	< .01	.95	.93	90.	.0407	.04	.62	192.68	I	I	I	I	I		I
Configural (non invariance)	256.42 106 < .01	106	< .01	.94	.92	.05	.0406	.05	.65	404.42	I	I	I	I	I		I
Metric (À equal)	294.00 118 < .01	118	< .01	.93	.92	.05	.0406	.06	.67	418.00	1	37.58	12	< .01	.01	0	01
Strong (t equal)	373.97 130 < .01	130	< .01	<u> </u>	<u> 06</u> .	.05	.0406	.06	.79	473.97	7	79.97	12	< .01	.03	0	0
Partial strong (tagg1, agg2)	323.44 126 <.01	126	<.01	.92	.92	.05	.0406	.05	69.	431.44	0	29.44	8	<.01	.01	0	0
Partial strict (ôagg1, agg2)	360.26 134 <.01	134	<.01	.92	.91	.05	.0406	90.	.73	452.26	e	36.82	8	<.01	.01	0	01

Approximation; 90% CI: Confidence Interval 90% for RMSEA; SRMR: Standardized Root Mean Square Residual; ECVI: Expected Cross-Validation Index; AIC: Akaike's *Notes.t*: Intercepts; χ^2 : Maximum Likelihood chi-square; df: degrees of freedom; TLI: Tucker-Lewis index; CFI: Comparative Fit Index; RMSEA: Root Mean Square Error of Information Criterion; Δx^2 : Difference in x^2 ; Δx^2 (ML): Difference in x^2 (ML); Δdll : Difference in degrees of freedom; ΔCFI : Difference in comparative fit index; $\Delta RMSEA$: Difference in mean square error of approximation; ΔSRMR: Difference in standardized root mean square residual. 683 684 685 685 686

687 Table III: Les corrélations du test-retest (4 semaines) pour les échelles de la VFEACC.

Scales	Temporal stability	М	SD	α	ICC	ICC 95% CI	р
	T1	2.60	0.83	.74	.71	.60– .82	< .01
Anger	T2	2.40	0.80	.71	.68	.54– .79	< .01
A	T1	2.04	0.75	.72	.70	.56– .80	< .01
Aggressiveness	T2	1.97	0.60	.69	.60	.55– .62	< .01
Total	T1	2.33	0.66	.79	.76	.56– .84	< .01
Total	T2	2.18	0.57	.70	.66	.51– .77	< .01

688 Table III: Four-week test–retest correlations for the FVCAAS subscales.

689 *Notes.* a: Cronbach's alpha; ICC: Intraclass Correlation Coefficient; ICC 95% CI: 95% Confidence Interval for 690 the ICC, n = 61.

⁵⁵⁰

691 Table IV: Liens de la VFEACC avec l'agression et le désengagement moral.

Scales	Anger ($\alpha = .78$)	Aggressiveness ($\alpha = .81$)	Total ($\alpha = .86$)
FR-BPAQ Physical ($\alpha = .75$)	.39*	.51*	.52*
FR-BPAQ Verbal ($\alpha = .80$)	.45*	.47*	.51*
FR-BPAQ Anger ($\alpha = .75$)	.57*	.29*	.50*
FR-BPAQ Hostility ($\alpha = .82$)	.42*	.38*	.45*
FR-BPAQ Total ($\alpha = .90$)	.52*	.46*	.57*
Moral disengagement ($\alpha = .88$)	.42*	.55*	.53*

692 Table IV: Relationships between the FVCAAS and aggression and moral disengagement.

693 *Notes.* *p < .05. FR-BPAQ: French version of Buss and Perry's Aggression Questionnaire, n = 175

694 Table V: Moyennes et écart-types des échelles de la FVCAAS en fonction du group.

Grouping factor	Group		Aı	nger	Aggress	siveness	Tot	al
		N	М	SD	М	SD	М	SD
All		175	2.45	0.81	1.95	0.80	2.20	0.71
	Males	106	2.58*	0.82	2.21*	0.81	2.93*	0.72
Gender	Females	69	2.25	0.75	1.55	0.59	1.90	0.59
	High-Contact	83	2.51	0.84	2.20*	0.89	2.35	0.77
Sport	Low-contact	92	2.40	0.78	1.71	0.63	2.06	0.63
	Males x High-Contact	63	2.61	0.76	2.36	0.85	2.48	0.27
Gender x Sport	Females x High-Contact	20	2.15	0.98	1.73	0.84	1.94	0.79
	Males x Low-contact	44	2.52	0.90	1.99	0.71	2.25	0.71
Gender x Sport	Females x Low-contact	49	2.29	0.64	1.47	0.44	1.88	0.49

695 Table V: Means and standard deviations for the FVCAAS scales organized by group.

696 Note. *p < .05, n = 175.

697Figure captions698Table1: Theoretical structure of the FVCAAS was similar to the original scale (CAAS)699Figure 1: Validity of the theoretical structure of the FVCAAS700Table 2: Factorial validity of the FVCAAS structure and its partial invariance across gender701Table 3: Temporal stability of the FVCAAS702Table 4: Theorical validity of the FVCAAS

Table 5: Usefulness of the FVCAAS to measur the differences amongst athletes' groups