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1 **Systematic Detection of Clustered Seismicity Beneath**
2 **the Southwestern Alps**

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8 **Key Points:**

- 9
- 10 • Automated earthquake detection combining beamforming, event classification and
11 template matching applied to the Southwestern Alps.
 - 12 • Systematic analysis of earthquake clustering by estimating the fractal dimension
13 of time series of earthquake occurrence.
 - 14 • Burst-like seismicity in fractured medium in the Briançonnais and Dora Maira vs
fluid-driven swarm-like seismicity in the Ubaye valley.

Abstract

We present a new automated earthquake detection and location method based on beamforming (or back projection) and template matching, and apply it to study the seismicity of the Southwestern Alps. We use beamforming with prior knowledge of the 3D variations of seismic velocities as a first detection run to search for earthquakes that are used as templates in a subsequent matched-filter search. Template matching allows us to detect low signal to noise ratio events, and thus to obtain a high spatiotemporal resolution of the seismicity in the Southwestern Alps. We describe how we address the problem of false positives in energy-based earthquake detection with supervised machine learning, and how to best leverage template matching to iteratively refine the templates and the detection. We detected 18,754 earthquakes over one year (our catalog is available online), and observed temporal clustering of the earthquake occurrence in several regions. This statistical study of the collective behavior of earthquakes provides insights into the mechanisms of earthquake occurrence. Based on our observations, we infer the mechanisms responsible for the seismic activity in three regions of interest: the Ubaye valley, the Briançonnais and the Dora Maira massif. Our conclusions point to the importance of fault interactions to explain the earthquake occurrence in the Briançonnais and the Dora Maira massif, whereas fluids seem to be the major driving mechanism in the Ubaye valley.

1 Introduction

Earthquake catalogs are the cornerstone of many studies in seismology, such as characterizing the seismic source (*e.g.* Abercrombie, 1995; Ide et al., 2003), estimating the amount of stress released at plate margins and understanding the role of repeating seismicity in this releasing process (*e.g.* Nadeau et al., 1995; Wech & Creager, 2011; Shelly et al., 2011; Frank et al., 2014), constructing reference earth models (*e.g.* Dziewonski & Anderson, 1981; B. Kennett & Engdahl, 1991; B. L. Kennett et al., 1995), seismic tomography (*e.g.* Dziewonski & Woodhouse, 1987; Van der Hilst et al., 1997; Li et al., 2008), seismic hazard estimation (*e.g.* on California Earthquake Probabilities, 1995), or modeling of the earthquake cycle (model calibration, *e.g.* Richards-Dinger & Dieterich, 2012). The first generation of regional and global catalogs were based on phase arrival picks on analog records (*e.g.* Engdahl et al., 1998). With the advent of digital recording, energy-based detection methods such as the short-term/long-term average (STA/LTA, Allen, 1982) method became popular.

The transition to digital recording and storage, the implementation of protocols for data curation and sharing, the increasing availability of data from networks and arrays, and the recognition of different types of earthquake signals motivated the development of more sophisticated earthquake detection and location algorithms, based, for instance, on array processing (*e.g.* Meng & Ben-Zion, 2017), or learning methods, such as neural networks (*e.g.* Perol et al., 2018). Automated data processing is not only essential for extracting signal from large, and rapidly increasing, data volumes, it also leads to uniform catalog quality.

Analysis of the seismic wavefield recorded at multiple sensors leverages the coherency of the signal across the station array to detect seismic phases which human eyes would have failed to identify. Network-based detection has led to the identification of phenomena such as low frequency earthquakes (*e.g.* Shelly et al., 2007; Brown et al., 2008; Frank et al., 2014) and non-volcanic tremor (*e.g.* Obara, 2002; Rogers & Dragert, 2003).

We develop an earthquake detection method that combines array processing, or, more precisely, a beamformed network response (Frank & Shapiro, 2014) and template matching (Gibbons & Ringdal, 2006; Shelly et al., 2007; Frank & Shapiro, 2014; Ross et al., 2019). Template matching is known to be efficient at detecting low signal-to-noise ratio (SNR) signals (*i.e.* with $\text{SNR} < 1$), and the required prior knowledge of the target seismicity is obtained from the beamformed network response.

67 We applied this new detection algorithm to one year of seismic data from 87 sta-
 68 tions located in the Southwestern Alps, between August 2012 and August 2013, includ-
 69 ing 55 stations from the temporary network CIFALPS (*cf.* Zhao et al., 2016, and see more
 70 information in Data and Resources). Although the Western Alps have been studied for
 71 a long time, the mechanisms driving the seismicity are still not well understood (*cf.* Noc-
 72 quet, 2012, and references therein), and a more complete earthquake catalog will make
 73 possible new studies to investigate the tectonic processes that cause them. The Alps were
 74 formed following the closure of the Alpine Tethys ocean, due to converging motion be-
 75 tween Europe and Africa. The mountain range is located at the border between the Eurasian
 76 plate and the Adriatic plate (*cf.* Figure 1). In the Western Alps, Chopin (1984) gave the
 77 first petrological evidence for continental subduction, which was later confirmed by sev-
 78 eral geophysical studies (*e.g.* Nicolas et al., 1990; Zhao et al., 2015). It is unclear, how-
 79 ever, whether subduction is still taking place. Even though geodetic data show that the
 80 Adriatic plate is rotating counterclockwise with respect to stable Europe (*e.g.* Serpel-
 81 loni et al., 2007), there is no observation of shortening in the Western Alps and part of
 82 the seismic activity is observed to occur under an extensional regime (*cf.* analysis of earth-
 83 quake focal mechanisms, Delacou et al., 2004). Various studies (*e.g.* Delacou et al., 2004;
 84 Nocquet et al., 2016; Walpersdorf et al., 2018) show that the earthquake activity in the
 85 Southwestern Alps is likely to be due to a complex combination of plate tectonic forces
 86 and other forces such as buoyancy forces or post glacial rebound. A more detailed char-
 87 acterization of seismic activity, which is indicative of active deformation, will help ad-
 88 dress these issues.

Figure 1. Interpretative cross-section of the Western Alps. Following the closure of the Alpine Tethys ocean, the collision of the European and Adriatic margins formed the Alps and the subduction complex illustrated here. A clear understanding of what is driving the deformation and the seismic activity in these complex geological units is still lacking. Abbreviations: FPF – Frontal Penninic Fault, Srp – serpentized, RMF – Rivoli-Marene deep fault. We show the locations of the CIFALPS stations on the topographic profile of the cross-section. The onset shows the location of the transect in the Western Alps, Europe. Figure modified from Zhao et al. (2015) and Solarino et al. (2018).

89 We first describe the earthquake detection method, and then present the earthquake
 90 catalog we thus obtained in the Southwestern Alps. We gain new insights into the seis-
 91 micity of the study region by investigating the collective behavior of earthquakes, made
 92 possible by the large number of detected events. We then discuss the importance of earth-
 93 quake interaction in the observed behavior of clustered seismicity.

94 2 Earthquake Detection Method

95 Detecting low SNR seismic signals by means of template matching requires knowl-
 96 edge of the type of signal to search for in the data. This can be obtained from an ex-
 97 isting earthquake catalog or from a preliminary detection run. Since the former is not
 98 publicly available for our study area, we produced a preliminary catalog using the energy-
 99 based detection method from Frank and Shapiro (2014), which is described in the fol-
 100 lowing. The events thus found were then used as template events in a subsequent matched-
 101 filter search.

102 2.1 Data Pre-processing

103 We used seismic data recorded between August 2012 and August 2013 at 87 seis-
 104 mic stations in the Southwestern Alps. The network includes 55 broadband sensors from

105 the temporary CIFALPS array (China-Italy-France Alps survey, Zhao et al., 2016, sam-
 106 pling at 100 Hz), and 32 broadband sensors from French and Italian networks (sampling
 107 at 100 Hz or 125 Hz, see Data and Resources). The data are downsampled to 50 Hz and
 108 filtered in the band 1-12Hz, which we found was a good compromise between targeting
 109 the frequency band of interest for observing local earthquakes and removing undesired
 110 signal.

111 2.2 Energy-based Detection (Composite Network Response)

112 The beamformed network response method due to Frank and Shapiro (2014) seeks
 113 to determine the origin, in time and space, of the seismic energy recorded at an array.
 114 This approach leverages the coherency of seismic energy across a receiver array for au-
 115 tomatic event detection. Using wave speeds according to a 3-D reference model (Potin,
 116 2016), the apparent travel times measured in the seismograms are then associated with
 117 a source location.

As a toy example, let us consider the earthquake whose location is indicated by a
 yellow star in Figure 2, and whose waveforms are recorded at multiple stations at the
 surface. Because spatial coherency of the seismic waveforms is not ensured (*e.g.* due to
 crustal heterogeneities or focal mechanism), we prefer to work with the envelopes of the
 waveforms. The envelope is the amplitude of the analytical representation of a time se-
 ries, it is calculated after the preprocessing described in Section 2.1 and the processing
 of the data is illustrated in Figure S1. We first discretize the volume beneath the study
 region into a grid of points, each of which representing a possible location of the seis-
 mic source (*cf.* Figure 2A). Each of these hypothetical sources is associated with a col-
 lection of P- and S-wave travel times to each of the stations. For a sufficiently accurate
 velocity model, the travel times from the potential source closest to the real source will
 provide the best alignment with the envelopes of the seismic data (*cf.* Figure 2B). We
 define the stack of the shifted envelopes as the network response:

$$NR_k(t) = \sum_{s,c} f(u_{s,c}(t + \tau_{s,c}^k)). \quad (1)$$

118 In Equation 1, k identifies a potential source and s, c are the station and the com-
 119 ponent indexes, respectively. We use the S-wave travel times on the horizontal compo-
 120 nents and the P-wave travel times on the vertical component; $\tau_{s,c}^k$ is the travel time from
 121 potential source k to station s on component c . u is the data and f is some transfor-
 122 mation of the seismic waveforms. In our case f relates to the function "envelope" (see Sup-
 123plementary Material Figure S1). The source k^* that yields the largest network response
 124 is found by a grid search and represents a proxy of the real source location. Locating earth-
 125 quakes through such a grid search, that is, shifting and stacking seismic energy, is also
 126 known as back projection or migration (*e.g.* Ishii et al., 2005; Walker et al., 2005; Honda
 127 & Aoi, 2009), but the objective here is detection.

For earthquake detection purposes, the quantity of interest is the largest network
 response of the grid at each time step. We define the composite network response (CNR)
 as:

$$CNR(t) = \max_k \{NR_k(t)\} = NR_{k^*}. \quad (2)$$

128 The process of searching for NR_{k^*} , continuously in time, is illustrated in Figure S3.

Figure 2C shows an example of CNR from real data. We postprocess the CNR by
 removing the baseline – a curve connecting the local minima – to set the noise level to
 zero (which explains the negative values in CNR). The peaks of CNR that exceed a user-
 defined threshold are detections of events, and the source locations are given by the cor-
 responding k^* . We use the following time-dependent threshold:

$$\text{threshold}(t) = \text{median}(CNR)(t) + 10 \times \text{MAD}(CNR)(t), \quad (3)$$

Figure 2. Top left panel (A): Spatial discretization of the volume beneath the study region. Using a velocity model, each point of the grid is associated with a collection of source-receiver travel times. The grid points are called potential seismic sources. As an example, let us consider an earthquake with location shown by the yellow star, and recorded at multiple stations. **Right panel (B):** The envelopes of the earthquake waveforms are shifted using the travel times of a potential seismic source close to the real location (*yellow star*). The shifted envelopes are then stacked to calculate the network response (*green waveform*, cf. Equation 1). The resulting network response is intrinsically related to the potential seismic source from which the travel times were calculated: different potential seismic sources give different network responses. **Bottom panel (C):** Composite network response (cf. Equation 2) calculated over one day. We subtract a curve connecting the local minima of the CNR to set its baseline to zero. To adapt to variations in the level of noise, we use a time-dependent threshold: the value "median + 10 × MAD" is evaluated every 30 minutes and a linear interpolation makes the threshold varying continuously within each 30-minute bin. Using small bin sizes enables the threshold to adapt to locally noisy episodes, but at the risk of discarding actual events: a 30-minute bin size is a good compromise between the two. We perform the peak selection on a smoothed CNR and impose a minimum peak distance, which explains why some of the values above threshold are not selected.

129 where MAD stands for median absolute deviation. We evaluate median (CNR) + 10 ×
 130 MAD (CNR) in 30-minute bins and make a continuously varying threshold by linearly
 131 interpolating the values obtained every 30 minutes.

132 Each detection yields a so-called template event (located at k^*), and the template
 133 for that event is then built by extracting waveforms using the detection time, travel times
 134 from k^* to each of the stations considered in the template (in our case, the 20 stations
 135 that are closest to k^*), and a window length (we choose 8 seconds). For our application
 136 in Section 3, we considered potential sources 1 km apart on a regular 3D cartesian grid
 137 (to 80 km depth) beneath a geographic area from 5.5°-9.0°E in longitude and 43.5°-46.0°N
 138 in latitude. This 1 km spacing is a good compromise between computation time, array
 139 sizes and detection performances.

140 2.3 Classification of Seismic Signals

141 Before using a template in a matched-filter search it is important to verify that the
 142 signal is due to an earthquake, because the CNR can be influenced by non-earthquake
 143 signals, such as proximal noise sources, electronic noise, and by issues in the preprocess-
 144 ing. For this purpose, we conduct a signal classification step prior to template match-
 145 ing.

146 For automated analysis and signal classification we use supervised machine learn-
 147 ing: to discriminate earthquakes from non-earthquakes, an algorithm is trained on a rel-
 148 atively small set of examples classified by a human expert. Our algorithm computes a
 149 linear combination of the signal features to generate a scalar that is fed into the logis-
 150 tic function (bounds the output between 0 and 1), which gives the probability of being
 151 an earthquake. Therefore, our algorithm is a binary logistic classifier. More information
 152 on the structure of the classifier is provided in Figure 3. For each three-component record
 153 extracted from the 20 stations, we calculate five features:

- 154 1. the amplitude maximum,
- 155 2. the first three statistical moments of the distribution of the peaks of the waveform
 156 autocorrelation function: variance, skewness, and kurtosis,

Figure 3. Left panel (A): We randomly sample detections from the database of candidate template events and identify each channel as *earthquake* or *non-earthquake*. We attribute the label *earthquake* to the detections with more than nine channels identified as earthquakes (*non-earthquake* otherwise). This arbitrary choice can be tuned in order to select more or less low SNR earthquakes in the template database. **Right panel (B):** Structure of our binary logistic classifier. The signal features are first preprocessed by standardizing them (*i.e.* removing the mean and setting the standard deviation to one) and bounding them between -1 and 1 through the use of hyperbolic tangent. A linear combination of the preprocessed signal features generates a scalar, which is fed into the logistic function (also called sigmoid function). The resulting output is bounded between 0 and 1, and is interpreted as the probability of being an earthquake. An output greater than 0.5 means the detection is more likely to be an earthquake than a non-earthquake. This algorithm was built using the Python library Keras (Chollet et al., 2015).

157 3. the maximum of the moving kurtosis along the extracted time series,

158 for a total of 300 features per event detection. The amplitude maxima help identify strong
 159 signals, the maximum of the moving kurtosis is sensitive to seismic phase arrivals, and
 160 the statistical moments of the autocorrelation function discriminate spikes (with large
 161 kurtosis) from impulsive earthquake waveforms. These features are not dependent on
 162 the relative phase of the signals, which renders them insensitive to small source mislo-
 163 cation.

164 For our application, Section 3.1, we manually labeled the waveforms of 500 detec-
 165 tions as earthquakes or noise (any non-earthquake signals). We note that labeling the
 166 waveforms currently prevents the full automation of the method, but it has to be done
 167 only once. In the training dataset, a 60 channel (20 stations \times 3 components) template
 168 event is labeled as an earthquake if more than nine channels were individually identi-
 169 fied as earthquake waveforms by eyes. This somewhat arbitrary criterion is used to re-
 170 ject the low SNR earthquakes that would not be interesting for use as template events,
 171 or which are not identified as earthquakes with high confidence. For training the algo-
 172 rithm, we split the dataset into two independent sub datasets: the training dataset (75%
 173 of the detections) and the validation dataset (25% of the detections). Each of these datasets
 174 were then augmented by a factor 100 by shuffling the channels in the templates (the clas-
 175 sification output must not depend on the order in which the input features are given).
 176 While optimizing the classifier with gradient descent on the training dataset, we evalu-
 177 ated the error on the validation dataset and stopped optimizing when this error began
 178 to increase. This method, which is known as early stopping (*e.g.* Yao et al., 2007), im-
 179 plicitly regularizes the classifier by providing a criterion for stopping the training when
 180 further updating the parameters would only overfit the data. On average, for several ran-
 181 domly selected training and validation datasets, we had a training accuracy of 0.92 and
 182 a validation accuracy of 0.90. Eventually, the classification process outputs a database
 183 of template events to be used in template matching.

184 2.4 Template Matching

185 In seismology, we often approximate the Earth as a linear filter and write an earth-
 186 quake seismogram as the convolution of a source term with a propagation term and an
 187 instrument term:

$$u(r, t) = \underbrace{S(t)M(r; \xi)}_{\text{source}} * \underbrace{G(r, t; \xi)}_{\text{propagation}} * \underbrace{I(t)}_{\text{instrument}} . \quad (4)$$

188 In Equation 4, the source term is the product of the source time function S and
 189 the focal mechanism M that describes effects due to preferred directions in the rupture
 190 process (*e.g.* rupture on a fault plane). The propagation term G , the Green’s function,
 191 describes how the earth responds to an impulsive source for a given travel path. We in-
 192 clude site effects in the Green’s function. I represents how the recording device distorts
 193 actual ground motion. The receiver location is r and the source location is ξ . Equation 4
 194 shows that colocated earthquakes produce similar waveforms because of similar Green’s
 195 functions. Moreover, similarity is high when the source functions have the same shape
 196 (similar focal mechanisms and magnitudes). Template matching leverages this expected
 197 similarity to detect new events.

Figure 4. **Left panel (A):** The waveforms of a template event (*red waveforms*), on 12 sta-
 tions and each of the 3 components, match well the data (*blue waveforms*): a new earthquake
 is detected. The correlation coefficient (CC) is given on each channel. **Right panel (B):** Com-
 parison of the template waveform on one channel (*red waveform*) with the waveforms of a few
 detected events (*blue waveforms*).

Template matching consists of scanning continuous recordings in search for matches
 between data and the waveforms that constitute a template. This method has proven
 to be efficient at detecting events with low SNR (SNR < 1, *e.g.* Gibbons & Ringdal, 2006;
 Shelly et al., 2007; Frank et al., 2014; Ross et al., 2019). Formally, scanning the data means
 calculating the correlation coefficient between the template waveforms and the data, con-
 tinuously in time. We use the following definition of the average correlation coefficient:

$$\text{CC}(t) = \sum_{s,c} w_{s,c} \sum_{n=1}^N \frac{T_{s,c}(t_n)u_{s,c}(t+t_n+\tau_{s,c})}{\sqrt{\sum_{n=1}^N T_{s,c}^2(t_n) \sum_{n=1}^N u_{s,c}^2(t+t_n+\tau_{s,c})}}. \quad (5)$$

198 In Equation 5, N is the length of the template waveform, n is a temporal index, and $w_{s,c}$
 199 is the weight attributed to station s and component c . If all weights are equal, with $w_{s,c} =$
 200 $1/N_s N_c$ (with N_s, N_c being the number of stations and components), then it is equiv-
 201 alent to calculating the arithmetic mean. For station s and component c , $T_{s,c}$ is the wave-
 202 form template, $u_{s,c}$ the continuous data, and $\tau_{s,c}$ the moveout (or time shift) in $u_{s,c}$. The
 203 time t is the detection time, meaning that the template window starts at time $\tau_{s,c}$ af-
 204 ter the detection time. The template windows start four seconds before the S wave on
 205 the horizontal components and one second before the P wave on the vertical component.
 206 We note that Equation 5 assumes the mean of $T_{s,c}$ and $u_{s,c}$ within each sliding window
 207 of length N is zero. We have shown in previous work that this assumption is correct when
 208 the data are filtered such that the lower non-zero period in the data is shorter than the
 209 window length (*cf.* Data and Resources and Beaucé et al., 2017). In the application pre-
 210 sented in Section 3, template matching was done with a detection threshold of eight times
 211 the daily root mean square (RMS) of the correlation coefficient time series. This detec-
 212 tion threshold is more conservative than the commonly used threshold of $8 \times \text{MAD}$ (*e.g.*
 213 Shelly et al., 2007; Brown et al., 2008; Baratin et al., 2018, $8 \times \text{RMS} \approx 12 \times \text{MAD}$).

214 Evaluating the correlation coefficient over long periods of time, and for many tem-
 215 plates, requires high performance computing to do it within a reasonable amount of time.
 216 We use the software Fast Matched Filter (Beaucé et al., 2017), which is particularly quick
 217 when run on graphics processing units (GPUs). The scanning process is illustrated in
 218 Figure 4. In the application to data from the Southwestern Alps we use just over 1,400
 219 templates, a template duration of 8 s (with 50 samples per second), and one year of con-
 220 tinuous data from 87 3-component stations, and we evaluated $\text{CC}(t)$ every sample. Eight
 221 seconds is a good compromise between extracting a representative chunk of the target
 222 waveform, and a reasonable computation time. Running our codes simultaneously on 12
 223 nodes equipped with one Tesla K20m GPU each took 12 h. As expected, reading oper-
 224 ations (I/O) of data and templates is the most time consuming task.

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2.5 Second Generation Templates

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As illustrated in Figure 4, a matched-filter search provides us with many repetitions of the same target waveform. By stacking the waveforms of the detected events we can enhance the SNR in the template waveform, which decreases the unwanted correlation component of the CC between data and noise in the template, thus improving the quality of the detection, and allows the template events to be located better.

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Non-linear stacking, like the Nth-root stack or the phase-weighted stack, greatly improves the SNR with respect to the linear stack, but also distorts the target waveform because of their non-linear nature. Even if it does not enhance SNR as much as non-linear stacking, we prefer the Singular Value Decomposition-based Wiener Filter (SVDWF) because it does not distort the waveform. SVDWF is based on the association of spectral filtering (keeping a limited number of singular vectors from the singular value decomposition) and Wiener filtering, and was initially developed for processing noise correlation functions (Moreau et al., 2017). For each station and each component, the matrix of detected events is first denoised using SVDWF, and a new template waveform is then obtained by stacking the denoised waveforms. Figure S4 illustrates the performance of these different stacking strategies.

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Detection and location involve finding the optimal network response for a given f in Equation 1. For detection purposes, we prefer using the envelope for f , but for location purposes, we choose f to be the kurtosis-based transform presented in Figure 5A (from Baillard et al., 2014). This transform makes the signal more sensitive to seismic phase arrivals and, thus, biases the CNR towards finding the travel times that align well with the seismic phase arrivals. Performing this relocation process on the second generation template waveforms reduces the spatial spread of the potential sources that yield a large CNR (*cf.* Figure 5, more details in Appendix A).

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The second generation templates are used in a subsequent matched-filter search to detect more events. This process – new template generation and matched-filter search – can be iterated several times until the earthquake catalog does not show notable updates between two iterations. During successive iterations, we optimize the template database by regrouping template events with same location and similar waveforms (template events with locations closer than 20 km and with average waveform correlation coefficient greater than 0.8) to avoid redundant matched-filter searches.

Figure 5. Relocation of the second generation templates. **Top panel (A):** The denoised and stacked waveforms obtained from the SVDWF (*blue waveforms*) are transformed following Baillard et al. (2014) to get a signal that is sensitive to phase arrivals (*orange waveforms*). The arrival times predicted by the new location are shown by black and red bars for the P- and S-wave, respectively. **Bottom left panel (B):** The composite network response (*blue curve*) is calculated using the orange signal shown in **A**. The neighborhood of the maximum of the CNR is analyzed to build a weighting function (*red curve*, *cf.* Appendix A for details). This weighting function is used to calculate a weighted average of the distance to the best potential seismic source (*cf.* Equation A3 in Appendix A), *i.e.* the potential source associated with the highest CNR. We define this weighted average as the uncertainty on the location. **Bottom right panel (C):** Each sample of the CNR shown in **B** is associated with a potential source in the grid; the color codes for the value of the CNR and the transparent points are those for which the weighting function is zero. In this example, the location uncertainty is 3.05 km.

3 Seismicity of the Southwestern Alps

We applied the earthquake detection method presented in Section 2 – that is, the combination of the Composite Network Response (CNR), signal classification, and template matching (with SVDWF) – to the preprocessed seismic data described in Section 2.1.

3.1 Catalog

Calculating the CNR as described in Section 2.2 yielded a total of 50,262 detections (candidate template events). After applying the classifier described in Section 2.3, we were left with 1,725 template events. We further reduced this number to 1,406 by re-grouping redundant template events (*cf.* Section 2.5); Figure 6 shows their locations. The matched-filter search yielded 18,754 non-redundant detections, with redundancy defined as events with similar waveforms (average CC > 0.8), detected within a time interval of three seconds and from template earthquakes located within 20 km from each other. This arbitrary choice may remove actual earthquakes from the catalog and leave some double counted events but produces a reasonable number of detected events. Our earthquake catalog is available online (see Data and Resources).

Figure 6. Locations of the 1,406 template events. Template events relocated with an uncertainty $\Delta r < 15$ km are shown with filled dots, and template events for which we did not find a reliable location are shown with open diamonds; the color scale codes for the depth of the events. Black inverted triangles are the seismic stations used in this study. We note that the uncertainty estimation described in Section 2.5 does not always perform well for deep events, which do not only feature simple P- and S-wave arrivals as assumed in the calculation of the network response. Therefore, a few events with $\Delta r < 15$ km still show odd locations (*e.g.* deep events located out of the group of deep earthquakes around Torino). The purple star indicates the epicenter of a $M_L 3.9$ earthquake that occurred in early October 2012, and which is important for the discussion in Section 4. The onset shows the position of the Western Alps in Europe. The *black dashed line* corresponds to the axis along which the stations from the CIFALPS network are deployed; this axis is used to project the locations of the template events for 2D cross sections.

To evaluate how well our detection method performs, we compared our catalog to the SISmalp catalog of Potin (2016). The number of events detected and located by our algorithm is more than an order of magnitude larger than the approximately 1,200 included in the SISmalp catalog for our study region; more details on the comparison with this catalog are given in Figures S5 and S6. The events that we seem to have missed all have magnitude less than one and most less than 0.4 (*cf.* Figure S7), which might explain inconsistencies in reported location or non-detection. We note here that other catalogs are also publicly available for this region, such as the Réseau National de Surveillance Sismique catalog with 383 events, and the Istituto Nazionale di Geofisica e Vulcanologia catalog with 743 events.

The temporal distribution of the 18,754 events is shown in Figure 7A with the daily seismic rate. We also report the magnitude of the events for earthquakes with $M > 1$ and located with high confidence ($\Delta r < 5$ km). These local magnitudes are based on waveform amplitude ratios, they were estimated following the procedure described in Appendix B. Amplitude ratios of events with $M < 1$ are contaminated by noise and therefore the resulting magnitude estimates are not meaningful. $M = 1$ is also where we observe the Gutenberg-Richter relation to break down (see Figure S8). The daily seismic rate shows continuous seismic activity in the Southwestern Alps, and reveals the existence of episodes of strong, burst-like seismicity (*e.g.* October 2012 and January 2013). Figure 7B shows the earthquake temporal distribution on recurrence time versus detec-

292 tion time graphs for three templates in distinct geographical regions: the Ubaye valley,
 293 the Briançonnais and the Dora Maira massif (*cf.* locations in Figure 1). The recurrence
 294 time is the time interval between two colocated earthquakes, and thus is defined template-
 295 wise. These three templates offer a representative view of the diversity of seismic behav-
 296 iors observed in our study region. The Ubaye valley hosts continuous seismic activity with-
 297 out clear sequences of foreshocks-mainshock-aftershocks, but the seismicity of the Briançonnais
 298 and the Dora Maira massif are dominated by burst-like episodes. These episodes are char-
 299 acterized by recurrence times spanning many orders of magnitudes, which is the signa-
 300 ture of temporal clustering. Seismicity in the Ubaye valley also differs from the burst-
 301 like seismicity observed in the Briançonnais and the Dora Maira massif by the smaller
 302 magnitude range it spans (*cf.* Figure 7B).

Figure 7. Left panel (A): Daily seismic rate (left axis, *blue continuous curve*) and daily
 magnitude distribution (right axis, *red dots*). Details on the local magnitude scale are given in
 Appendix B. **Right panels (B):** Recurrence time vs detection time for three templates located
 in three distinct geographic regions. The Briançonnais and the Dora Maira massif are dominated
 by episodes of burst-like seismicity, and the Ubaye valley hosts continuous seismic activity that
 does not feature clear foreshocks-mainshock-aftershocks sequences. Local magnitudes are coded
 in color: we observe a smaller magnitude range in the Ubaye valley than for the earthquake
 sequences in the Briançonnais and in the Dora Maira massif.

303 3.2 Temporal Clustering of the Seismicity

304 Unlike Poisson seismicity, clustered earthquake sequences have earthquake occur-
 305 rence that is not random in time: instead, time clustered seismicity suggests that past
 306 events influence the occurrence of future ones. We emphasize that an earthquake sequence
 307 with high seismic rate does not have to be clustered in time, but can be Poissonian (*e.g.*
 308 Frank et al., 2018). Temporal clustering is often observed for sequences of foreshocks-
 309 mainshock-aftershocks (*e.g.* Utsu, 1961; Knopoff, 1964; Gardner & Knopoff, 1974; Zali-
 310 apiin & Ben-Zion, 2013a) and is thought to be the signature of stress redistribution on
 311 neighboring faults taking place during the seismic rupture (*e.g.* Burridge & Knopoff, 1967;
 312 Dieterich, 1992; Stein, 1999). More generally, temporal clustering can be explained by
 313 various mechanisms implying interactions between earthquakes (*e.g.* Frank et al., 2016).
 314 The observation of temporal clustering thus provides a window into the mechanisms of
 315 earthquake occurrence.

316 Quantifying the degree of temporal clustering requires characterization of the time
 317 series of earthquake occurrence. While accurate knowledge of the earthquake locations
 318 and magnitudes allows sophisticated characterization of clustering in the time-space-energy
 319 domain (*e.g.* Zaliapin et al., 2008; Zaliapin & Ben-Zion, 2013b), restricting the analy-
 320 sis to the time-space domain is an appropriate choice for the Southwestern Alps since
 321 earthquake magnitudes are small. To describe seismic activity, we introduce the event
 322 count $e(t)$ (*cf.* Figure 8A), that is, the number of events in narrow time windows (bins).
 323 We characterize clustering by means of the autocorrelation and spectrum of $e(t)$ (Fig-
 324 ure 8B and C). By definition, temporal clustering implies temporal correlation of the earth-
 325 quake occurrence at non-zero correlation time in the autocorrelation function. We ob-
 326 serve that clustered earthquake sequences exhibit power-law dependence of $e(t)$ on fre-
 327 quency ($\tilde{e}(f) \propto f^{-\beta}$, similar to Frank et al., 2016). The strength of temporal cluster-
 328 ing is quantified by β , referred to as clustering coefficient, which can be estimated from
 329 the slope of the spectrum in log-log space (Figure 8C). A strongly clustered earthquake
 330 sequence has a large β whereas an earthquake sequence close to a Poisson sequence has
 331 a small β , and $\beta = 0$ indicates a purely random sequence (flat spectrum).

Processes exhibiting a power-law spectrum are scale-invariant processes, within a certain range of scales limited by natural bounds. For instance, we expect the power-law $\tilde{e}(f) \propto f^{-\beta}$ to hold between the period of activation of the fault/seismic source (smallest frequency) and the smallest time interval we can resolve between two earthquakes (highest frequency). A powerful analysis tool for scale-invariant time series comes from the theory of fractal clustering (*e.g.* Turcotte, 1997; Lowen & Teich, 2005). Fractal analysis, which has been applied to earthquake occurrence in various studies (*e.g.* Smalley Jr et al., 1987; Lee & Schwarcz, 1995), consists of counting earthquakes in time intervals of variable width. In the case of fractal clustering, the fraction of occupied intervals x has a power-law dependence on the size of the intervals τ , *i.e.* $x \propto \tau^{1-D}$. The fractal dimension D is zero for a Poisson distributed earthquake occurrence, and is typically larger than 0.2 for clustered seismicity (*cf.* Figure 8D). We used correlation time, clustering coefficient β and fractal dimension D to characterize the temporal clustering in our study region. We found that the clustering coefficient was well appropriate for studying clustering over short times, whereas the fractal dimension gave the most contrasted results for studying the long-term clustering (see Supplementary Material Figure S9 and Figure S10). We present our observations of temporal clustering in Figure 9.

Figure 8. Quantification of temporal clustering. **Top left panel (A):** Event count number $e(t)$ for earthquakes detected with two different templates. The event count number is calculated by dividing the time axis into 5-minute bins, and counting the number of events within each bin. **Top right panel (B):** Autocorrelation function of the event count number. We define the correlation time τ as the time interval over which the autocorrelation function is greater than the threshold plotted with the dashed black line (arbitrarily set to 0.12). **Bottom left panel (C):** Power spectral density of the event count number. The spectrum of the event count number has a power-law dependence on the frequency when temporal clustering occurs. We define the power-law exponent β as the clustering coefficient. **Bottom right panel (D):** Fractal analysis of the earthquake sequences. Within a limited range of size of time intervals, the fraction of occupied intervals follows a power-law, whose exponent is related to the fractal dimension of the earthquake occurrence.

Comparison between Figure 9A and Figure 9B shows that there is no trivial correlation between the number of earthquakes per template (*i.e.* number of earthquakes in some volume around the template location) and temporal clustering. We distinguish three geographic regions of high seismic activity: from west to east, the Ubaye valley, the Briançonnais and the Dora Maira massif. The largest temporal clustering is observed beneath the western part of the Dora Maira massif (*cf.* the geological cross-section in Figure 1). The fractal dimension of the event count reveals large temporal clustering also in the southwestern part of the Briançonnais (fractal dimension $D \gtrsim 0.2$). Although we detected a large number of earthquakes beneath the Ubaye valley, we do not observe significant temporal clustering. The seismic activity in the Ubaye valley features a mixture of continuous unclustered seismicity punctuated by episodes of strong, clustered seismicity (see Supplementary Material Figure S9). The Ubaye valley is known to host a seismic swarm (*e.g.* Jenatton et al., 2007; Daniel et al., 2011; Leclère et al., 2012, 2013) that was reactivated in February 2012 by a M3.9 earthquake (Thouvenot et al., 2016). In the following discussion, we refer to swarms as episodes of high seismic activity without substantial temporal clustering (as in, for example, Zaliapin & Ben-Zion, 2013a).

4 Discussion

Frank et al. (2016) present a model where a group of stationary Poisson point processes can lead to a clustered event occurrence if there is interaction between the point

368 processes. They show that without interaction a coherent acceleration of the Poisson event
 369 rates cannot reproduce the clustered distribution as in Figure 8. Poisson point processes
 370 describe earthquake occurrences on faults experiencing constant tectonic loading. There-
 371 fore, temporal clustering is the signature of earthquake interaction rather than an increase
 372 of the external forcing of the faults (*e.g.* because of aseismic slip occurring in the vicini-
 373 ty). We note that elastic interactions are commonly invoked to explain time clustered
 374 events (*e.g.* Knopoff, 1964; Dieterich, 1994; Stein, 1999). Thus, assuming there exists
 375 a constant loading acting on the faults, we can expect systems with many interacting
 376 elements – dense fault networks or single faults with many asperities – to be able to pro-
 377 duce strong short-term clustering whereas clustering in sparser networks takes place on
 378 longer time scales. With our observations we cannot differentiate between multiple faults
 379 or single faults with multiple asperities at the sub-template scale (*i.e.* for events detected
 380 with the same template).

Figure 9. Cross-section along the CIFALPS axis showing 976 templates that were well re-
 located ($\Delta r < 15$ km). **Top panel (A):** Number of detected earthquakes per template. **Bottom
 panel (B):** Sources with fractal dimension $D > 0.2$, *i.e.* sources exhibiting temporal clustering.
 The fractal dimension was calculated by taking the event count $e(t)$ of each template plus all
 the templates within a 10-km radius, over the whole study period. Even though intense seismic
 activity is located in the Ubaye valley, this seismicity is not associated with significant temporal
 clustering, showing that there is no systematic relation between temporal clustering and number
 of events per unit volume. The purple star indicates the location of the $M_L 3.9$ earthquake that
 we mention in the discussion (Section 4). The red structures are reported from the geological
 cross-section in Figure 1.

381 Both regions where we observe significant temporal clustering, the Briançonnais
 382 and the Dora Maira massif, seem to share a common mechanism for clustering. Solarino
 383 et al. (2018) observed high V_p/V_s ratios (low V_s) in the uppermost part of the Briançonnais,
 384 where we observe high temporal clustering. They suggested that low shear wave veloc-
 385 ities V_s could be explained by the widespread fault network observed in the Briançonnais
 386 (*e.g.* Tricart et al., 2004). In the Dora Maira massif, all the templates detecting seismic
 387 activity with fractal dimension $D > 0.3$ (see Figure 9B) are located around the $M_L 3.9$
 388 earthquake that occurred on October 3rd, 2012 (*cf.* location in Figures 6 and 9, *cf.* our
 389 catalog for the local magnitude). This highly clustered seismicity took place over about
 390 four days (see Figure 7A), and can be seen as a sequence of foreshocks-mainshock-aftershocks.
 391 The locations shown in Figure 9 are substantially spreaded, which suggests that seismic-
 392 ity is occurring on multiple faults. Given the limited temporal extent of the episode, we
 393 expect fault interactions to be a major contribution to temporal clustering in this area.
 394 Moreover, it is known that the Dora Maira massif is made of ultra-high pressure meta-
 395 morphic rocks, *i.e.* of European crust subducted to 90 km depth and later exhumed (Chopin,
 396 1984), it is very likely to be fractured. Thus, along with geological evidence, our obser-
 397 vations of temporal clustering support the idea of fault interactions in dense fault net-
 398 works as a driving mechanism for clustering in the Briançonnais and the Dora Maira mas-
 399 sif.

400 Despite the high density of seismic sources beneath the Ubaye valley, temporal clus-
 401 tering is limited (only a few templates with $D \gtrsim 0.2$), which is an expected feature for
 402 seismic swarms. Thus, our measurements of temporal clustering suggest that the driv-
 403 ing mechanism for seismicity in the Ubaye swarm differs from the one in the Briançonnais
 404 and the Dora Maira massif. Multiple studies (*e.g.* Daniel et al., 2011; Leclère et al., 2012;
 405 De Barros et al., 2019) emphasized the role of fluids in the stressing mechanism driving
 406 the seismicity of the Ubaye swarm. Furthermore, Ben-Zion and Lyakhovskiy (2006) stud-
 407 ied numerically the influence of damage rheology on the production of earthquakes. Their

408 model shows that cold, brittle media produce burst-like seismicity (high temporal clus-
 409 tering) whereas regions with high fluid activity produce more diffuse, swarm-like seis-
 410 micity (low temporal clustering). Our observations of high seismic activity with low tem-
 411 poral clustering in the Ubaye swarm thus support the important role of fluid activity in
 412 this region. We realize, however, that such swarm-like behavior could also be the signa-
 413 ture of aseismic processes (*e.g.* Lohman & McGuire, 2007). Whether aseismic slip is an
 414 important factor (Leclère et al., 2013) or not (De Barros et al., 2019) is still an ongo-
 415 ing debate, and our observations are not enough to support one scenario over the other.
 416 The clustered seismicity we detected in the Ubaye valley is consistent with the obser-
 417 vations in De Barros et al. (2019) of coexisting aftershock sequences and swarm-like seis-
 418 micity in this area. Studying a longer period of time, including the 2003-2004 and 2012-
 419 2015 Ubaye seismicity, could provide information on the stationarity of temporal clus-
 420 tering in the Ubaye valley and the rest of the Southwestern Alps.

421 5 Conclusion

422 In this paper we present a new method for automated earthquake detection and
 423 location, based on template matching and beamforming (or back projection), and use
 424 it for high (spatiotemporal) resolution characterization of seismicity in the Southwest-
 425 ern Alps. We address the problem of false positives in energy-based detection with sig-
 426 nal classification based on supervised machine learning (Section 2.3), and we construct
 427 low noise templates by combining the singular value decomposition Wiener filter (SVDWF)
 428 with subsequent stacking (Section 2.5).

429 In our application to data from CIFALPS (Zhao et al., 2016), a semi-linear seis-
 430 mic network, and other permanent seismic stations in the Southwestern Alps, we detected
 431 in one year over an order of magnitude more events (18,754 vs. approximately 1,200) than
 432 an existing catalog based on traditional phase picking. We analyzed the statistical prop-
 433 erties of the seismicity, and observed and characterized temporal earthquake clustering.
 434 We observed that regions of high seismic activity and high temporal clustering coincided
 435 with regions that are highly fractured (Briançonnais) or likely to be fractured (Dora Maira
 436 massif). Seismicity in the Dora Maira massif during the study period was dominated by
 437 the sequence of foreshocks and aftershocks associated with the 2012-10-03 M_L 3.9 earth-
 438 quake. We also identified one region of high seismic activity and low temporal cluster-
 439 ing coinciding with the Ubaye swarm. Our results support interpretations invoking an
 440 important role of fluids in swarm seismicity (Daniel et al., 2011; Leclère et al., 2012, 2013;
 441 De Barros et al., 2019).

442 The efficiency of this method increases when the database of templates gets more
 443 complete. Thus, processing longer times is likely to give better results as the opportu-
 444 nities of detecting new template events grow. The systematic application of this method
 445 to the Western Alps data, or even to the whole mountain range, will help gathering new
 446 observations of the seismicity and understanding the tectonic context of the region. We
 447 note that even though we presented an application to a semi-linear seismic network, our
 448 method can be applied to any network geometry. If 3D wave speed variations are suf-
 449 ficiently well known on the scale of study, it is possible to perform comprehensive stud-
 450 ies of 3D seismicity structures by applying this method with 2D seismic arrays.

451 Data and Resources

452 The timings and locations of the 18,754 earthquakes we detected are available at
 453 E. Beaucé’s personal website <https://ebeauce.github.io/> in the Material section. The
 454 reported times are the origin times, so that users can retrieve the P- and S-wave data
 455 by adding the origin times and the travel times, also provided in the catalog. Our codes
 456 are available at https://github.com/ebeauce/earthquake_detection_EB_et_al_2019
 457 (last accessed 08/16/2019), and are provided with a real-data example.

We created the map in Figure 6 using the topographic data from the Shuttle Radar Topographic Mission (SRTM) 90m database (<http://www.cgiar-csi.org/data/srtm-90m-digital-elevationdatabase-v4-1>, last accessed May 2019). Our data come from the temporary experiment CIFALPS (Zhao et al. (2016), DOI: <http://dx.doi.org/10.15778/RESIF.YP2012>) and permanent French (FR and RD RESIF (1995)) and Italian (GU University of Genova (1967), IV INGV Seismological Data Centre (2006), MN MedNet Project Partner Institutions (1990) and MT OGS (Istituto Nazionale di Oceanografia e di Geofisica Sperimentale) and University of Trieste (2002)) networks. The RENASS and INGV catalogs we mention in Section 3.1 can be obtained at <https://renass.unistra.fr/recherche> and <http://cnt.rm.ingv.it/en>, respectively.

Our study showing that the simplified definition of the correlation coefficient we use in this work is valid is available at https://github.com/beridel/fast_matched_filter/blob/master/consequences_nonzero.pdf.

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The reader is referred to the Data and Resources section to find the origin of the data.

Appendix A Template Relocation

The weighting function presented in Figure 5 is defined by:

$$w(t_n) = \begin{cases} A \exp\left(-\frac{(\text{CNR}(t_n) - \text{CNR}_{\max})^2}{4\sigma}\right) & \text{if } t_n \in \mathcal{V}(t_{\max}), \\ 0 & \text{otherwise.} \end{cases} \quad (\text{A1})$$

In Equation A1, the neighborhood $\mathcal{V}(t_{\max})$ is defined by:

$$\mathcal{V}(t^*) = \{t^- \leq t_k \leq t^+ \mid [t^-, t^+] \text{ is a convex set, } t_{\max} \in [t^-, t^+], \text{CNR}(t_k) > 0.75 \times \text{CNR}_{\max}\}, \quad (\text{A2})$$

and $t_{\max} = \underset{t_n}{\operatorname{argmax}}(\text{CNR})$. A is a normalization factor such that $\sum_{n=1}^N w(t_n) = 1$, and σ is the standard deviation of the CNR within $\mathcal{V}(t_{\max})$.

Using the locations of the potential sources from the composite network response, we calculate the average distance to the best test source:

$$\Delta r = \sum_{n=1}^N w(t_n) |r_n - r_{\text{best}}|. \quad (\text{A3})$$

In Equation A3, N is the temporal length of the stacked waveforms, r_n is the potential source location associated with the CNR at time t_n and r_{best} is the location of

493 the potential source associated at time t_{\max} , *i.e.* the location of the second generation
494 template.

495 Appendix B Magnitude Estimation

496 Our local magnitude is calculated from the amplitude ratio of the peak velocities
497 with a reference event. Thus, we first need to estimate the magnitude of at least one event
498 per template to calibrate our local magnitude scale. For each family of earthquakes de-
499 tected with the same template, we proceed as follows:

- 500 1. calculate the S-wave spectrum on every station and component,
- 501 2. calculate the noise spectrum in a window taken just before the P-wave arrival,
- 502 3. average the spectra over all the stations and components, including only the sam-
503 ples satisfying the SNR criterion (similarly to Uchide & Imanishi, 2016), accord-
504 ing to:

$$\bar{S}(f) = \frac{1}{\sum_{s,c} \mathbf{1}_{\text{SNR}>5} [S_{s,c}(f)]} \sum_{s,c} \alpha_{s,c} S_{s,c}(f) \mathbf{1}_{\text{SNR}>5} [S_{s,c}(f)]. \quad (\text{B1})$$

505 In Equation B1, $\mathbf{1}_{\text{SNR}>5} [S(f)]$ is the indicator function testing whether $S(f)$ has
506 SNR greater than 5 (equal to 1) or not (equal to 0). The SNR is calculated at ev-
507 ery frequency by taking the ratio of the S-wave spectrum to the noise spectrum.
508 $\alpha_{s,c}$ is a corrective factor that we describe further.

- 509 4. The average spectra are converted to displacement spectra by using the relation-
510 ship

$$|u_{\text{velocity}}(f)| = f \times |u_{\text{displacement}}(f)|, \quad (\text{B2})$$

- 511 5. the average displacement spectra are fitted with the Boatwright model (Boatwright,
512 1978):

$$S_{\text{Boatwright}}(f) = \frac{\Omega_0}{\left(1 + \left(\frac{f}{f_c}\right)^4\right)^{1/2}}, \quad (\text{B3})$$

513 where Ω_0 is the low-frequency plateau, related to the seismic moment, and f_c is
514 the corner frequency.

515 The corrective factors $\alpha_{s,c}$ are defined such that the low-frequency plateau Ω_0 can
516 be identified to the seismic moment M_0 . Assuming a double-couple source, a dis-
517 placement amplitude spectrum can be written as (following Boatwright, 1978):

$$\begin{aligned} |u^S(f)| &= \frac{R^S}{2\rho\beta^3r} \frac{M_0}{\left(1 + \left(\frac{f}{f_c^S}\right)^4\right)^{1/2}} \exp\left(-\frac{\pi ft^S}{Q^S}\right), \\ \implies M_0 &= \Omega_0 \frac{2\rho\beta^3r}{R^S} \exp\left(\frac{\pi ft^S}{Q^S}\right), \\ \implies \alpha_{s,c} &= \frac{2\rho\beta^3r_{s,c}}{R^S} \exp\left(\frac{\pi ft_{s,c}^S}{Q_{s,c}^S}\right). \end{aligned} \quad (\text{B4})$$

518 In Equation B4, we use typical values for the S-wave velocity β (3000 km/s), the den-
519 sity ρ (2700 kg/m³) and the average S-wave radiation pattern R^S ($\sqrt{2/5}$ from Aki &
520 Richards, 2002). The seismic moment M_0 gives the magnitude moment M_w through:

$$M_w = \frac{2}{3} (\log M_0 - 9.1). \quad (\text{B5})$$

521 The reference events are those for which fitting a Boatwright model to the aver-
 522 age spectrum results in a variance reduction greater than 0.95. Figure B1 shows an ex-
 523 ample of an average spectrum that was fitted correctly, and therefore kept as a reference
 524 event. The local magnitude of all the other events are determined by:

$$M_i = M_{\text{ref}} + \text{Median}_{s,c} \left\{ \log \frac{A_{s,c}^i}{A_{s,c}^{\text{ref}}} \right\}, \quad (\text{B6})$$

525 or more generally if there are several reference events:

$$M_i = \text{Median}_k \left\{ M_{\text{ref},k} + \text{Median}_{s,c} \left\{ \log \frac{A_{s,c}^i}{A_{s,c}^{\text{ref},k}} \right\} \right\}. \quad (\text{B7})$$



Figure B1. Magnitude estimation of the reference event. For each template, we use the highest SNR detections to calculate the average S-wave spectrum (Equation B1) and fit it with the Boatwright model (Equation B3). The low-frequency plateau gives us the seismic moment M_0 . The average is calculated over all the stations and components that satisfy the SNR criterion. Thus, for each frequency sample the number of channels included in the average may vary, as we can see with the color scale. Since frequency samples with a higher number of channels are more reliable, we give them larger weight in the inversion.

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Figure 1.

CIFALPS stations

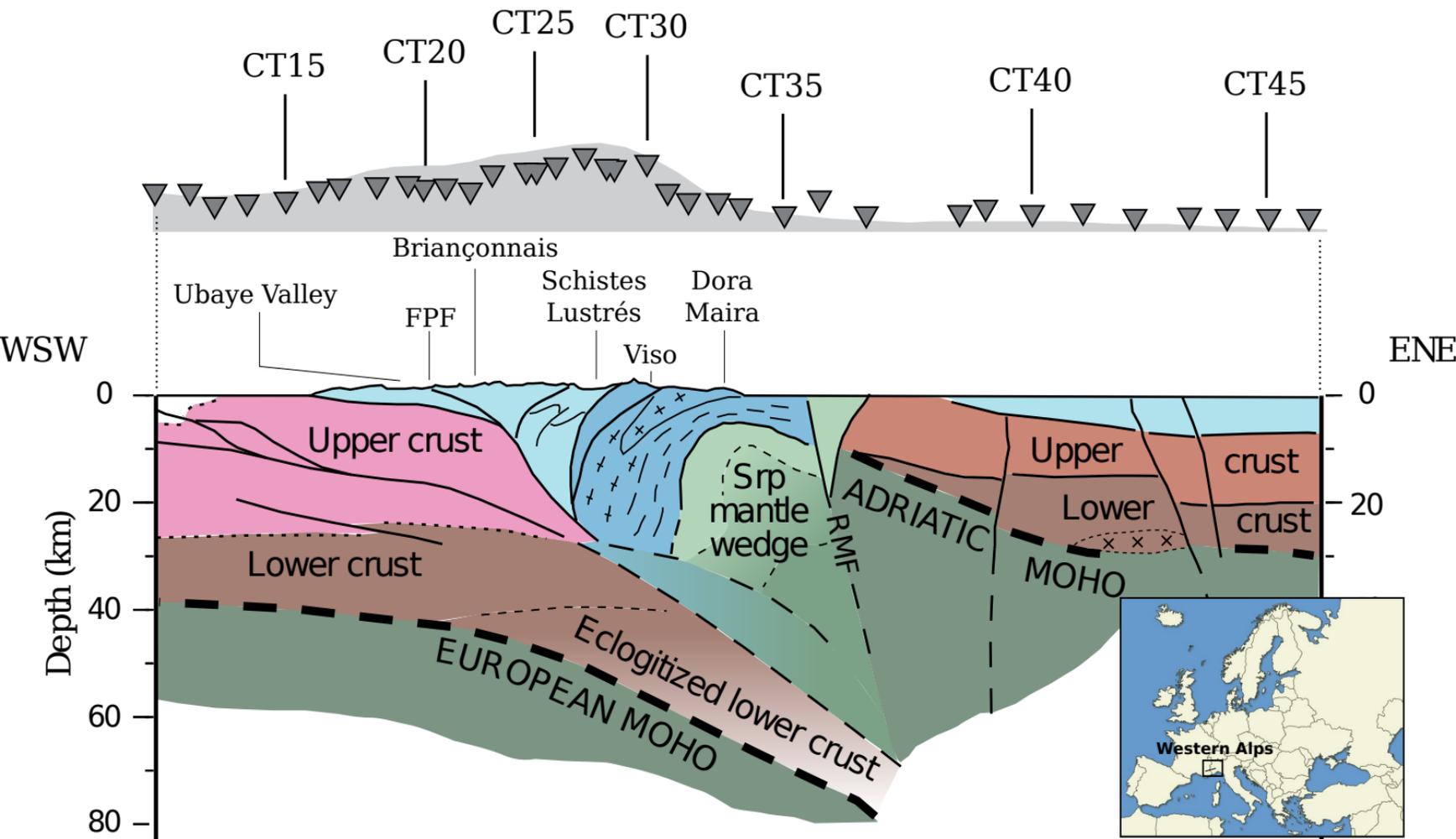


Figure 2.

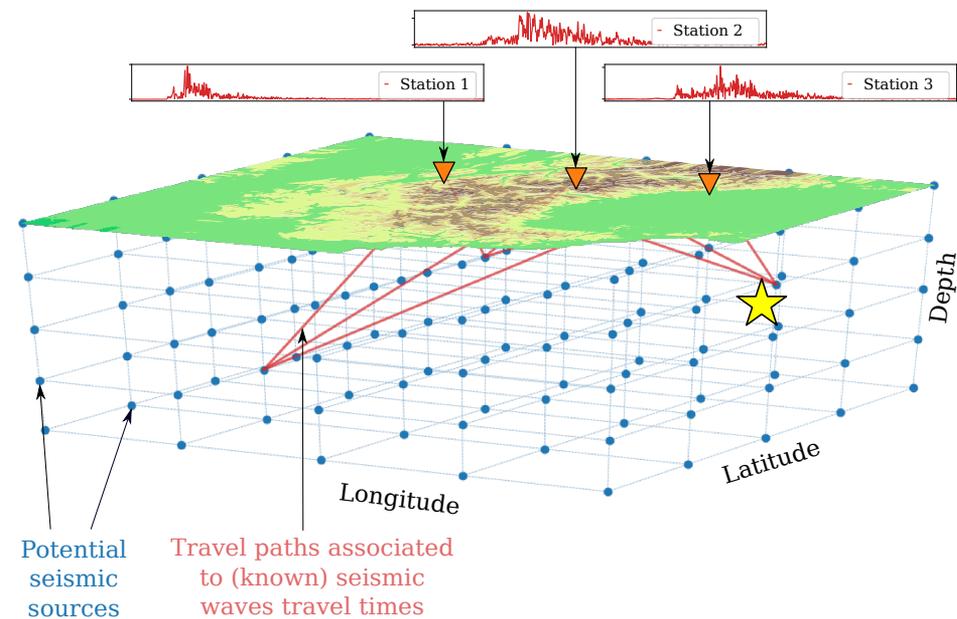
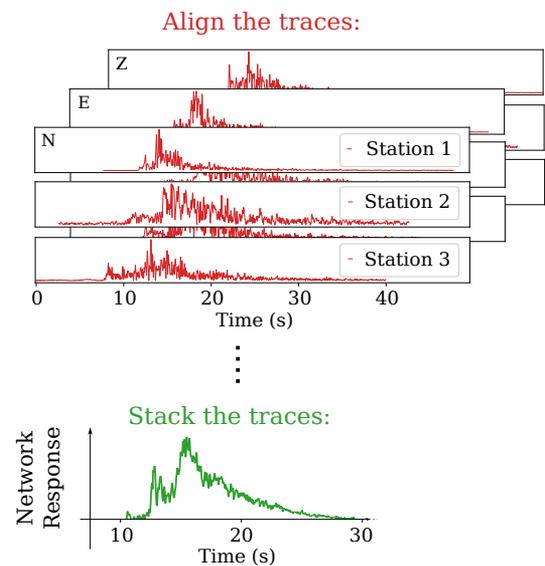
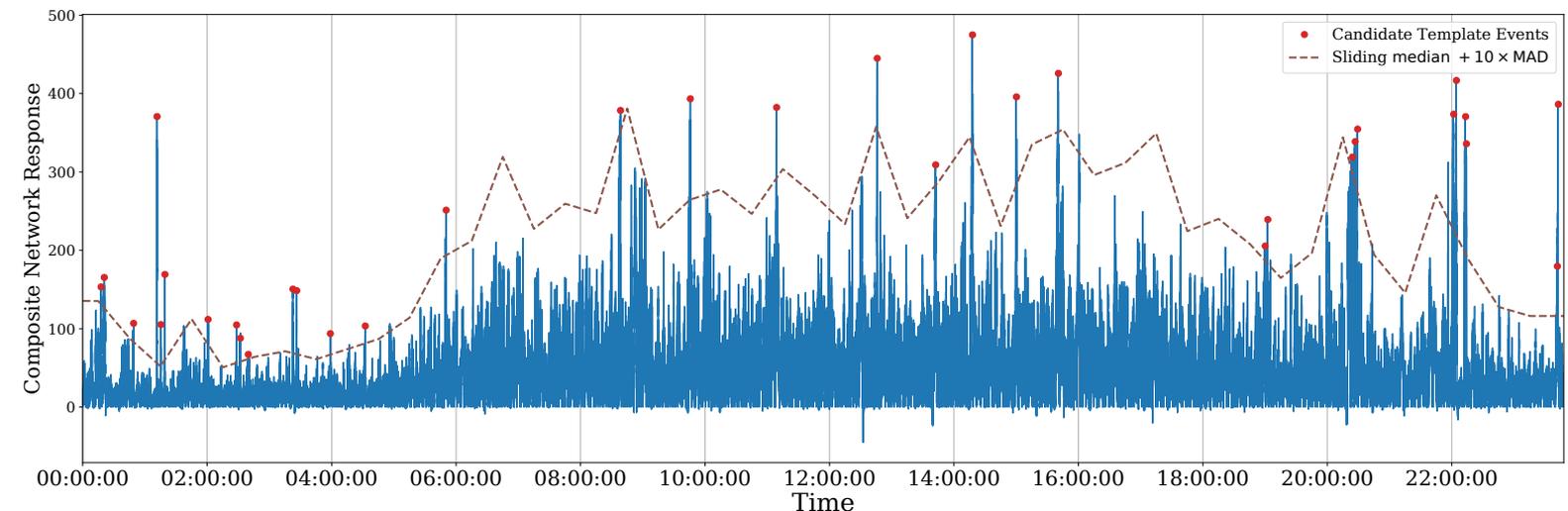
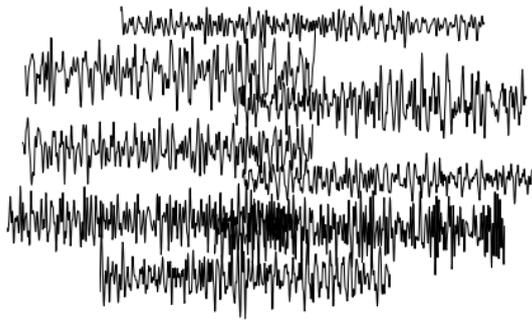
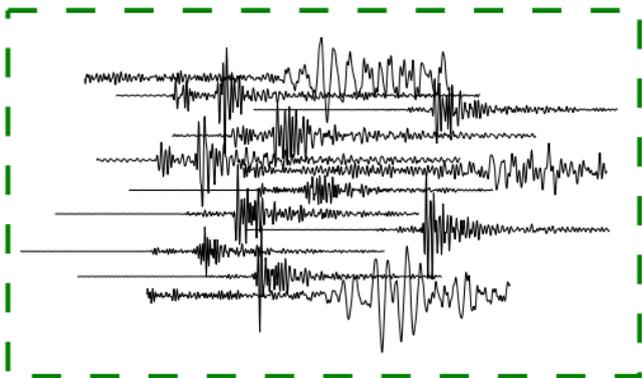
A**B****C**

Figure 3.

A**Seismic Noise****Earthquakes**

Training Dataset

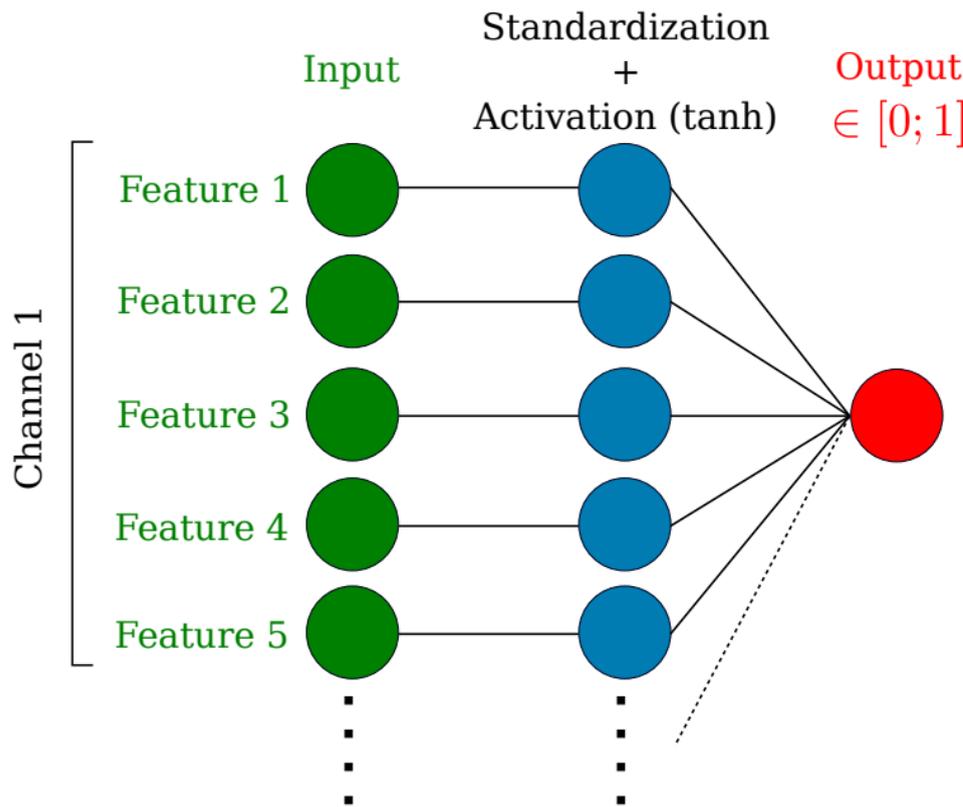
B

Figure 4.

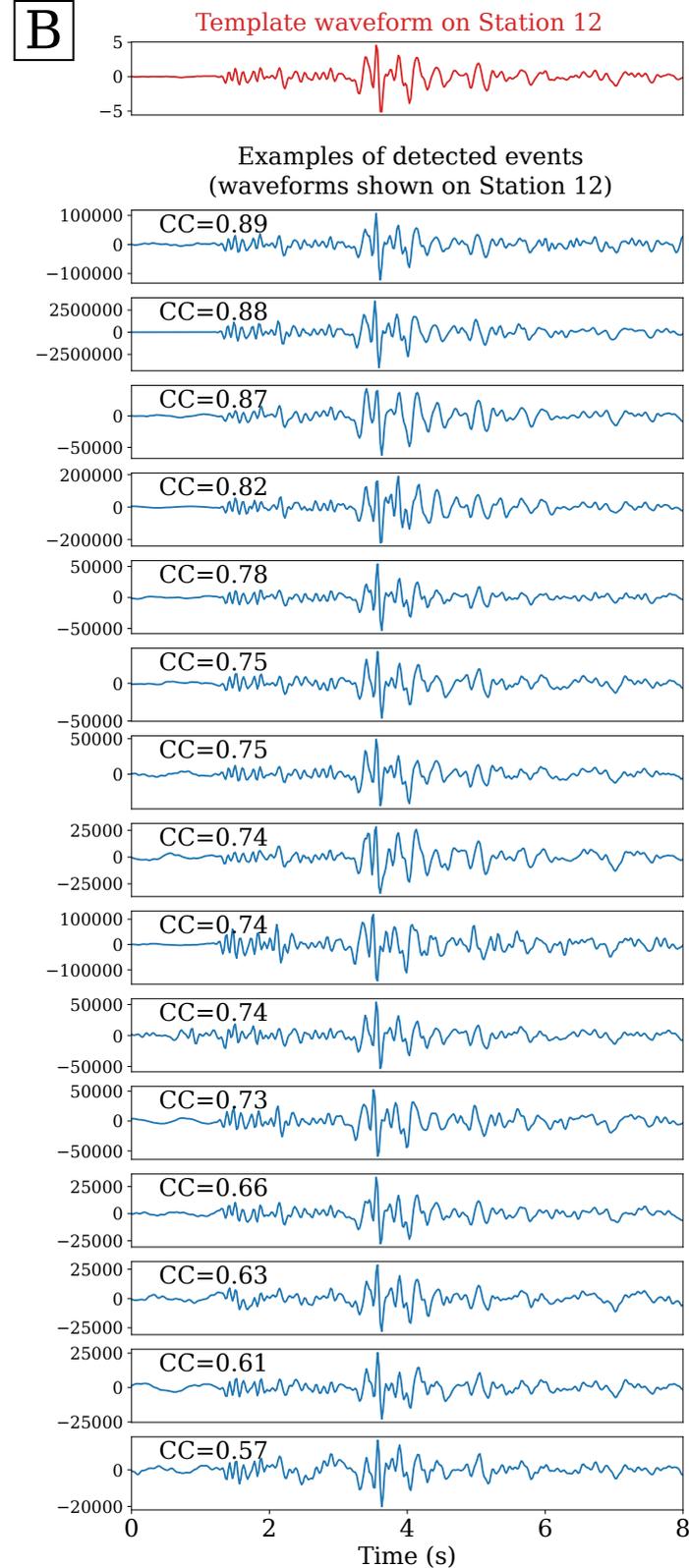
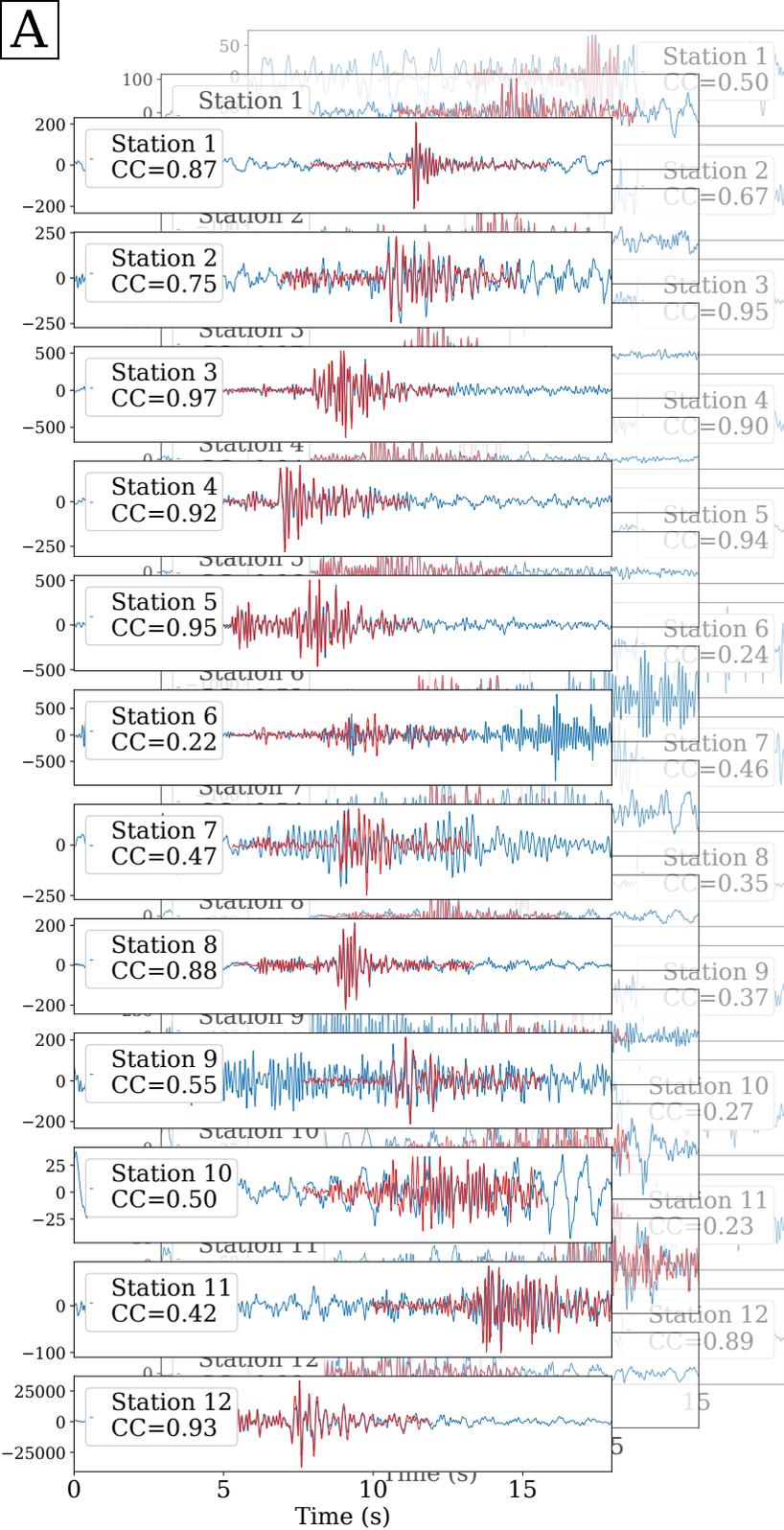


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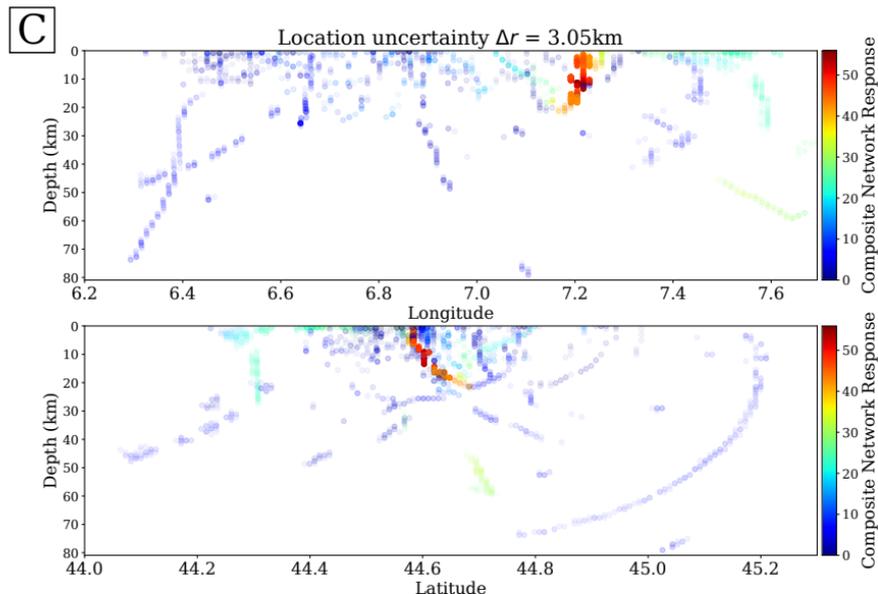
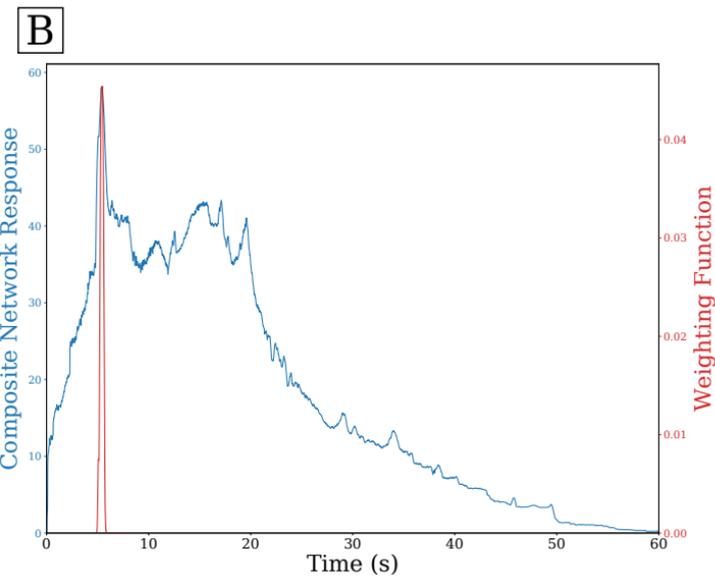
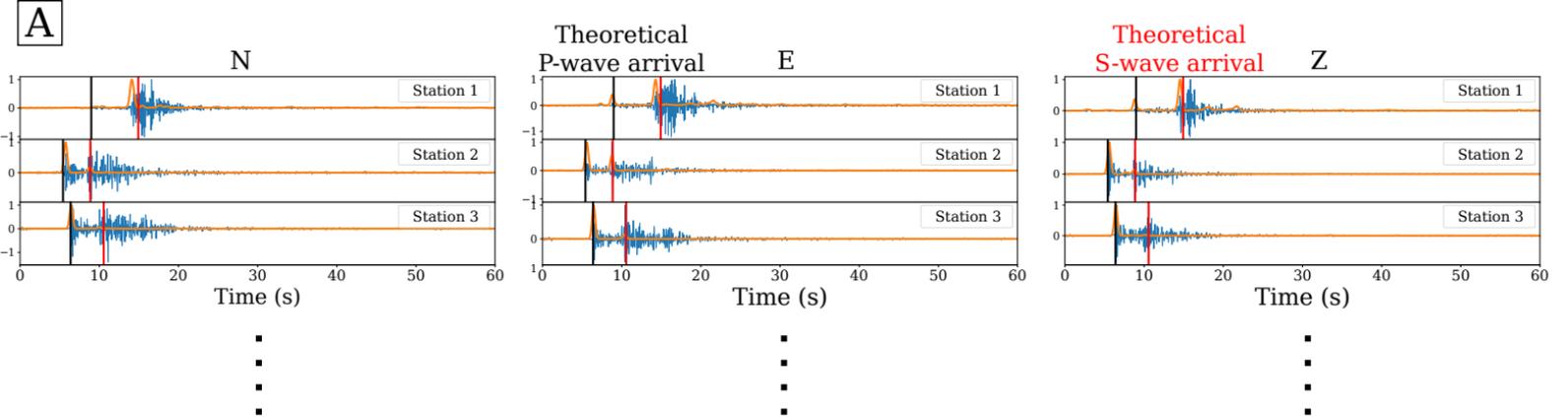


Figure 6.

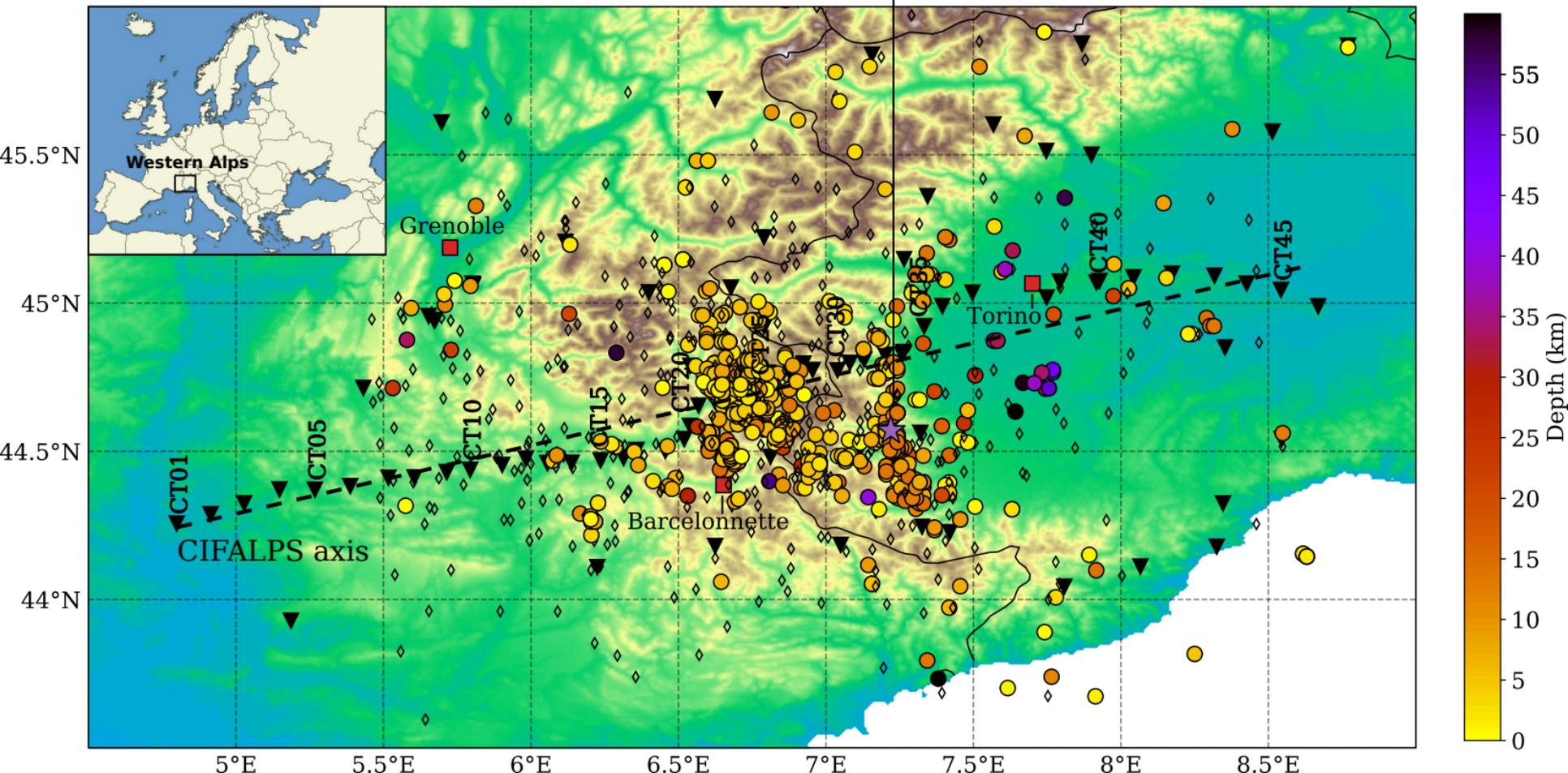


Figure 7.

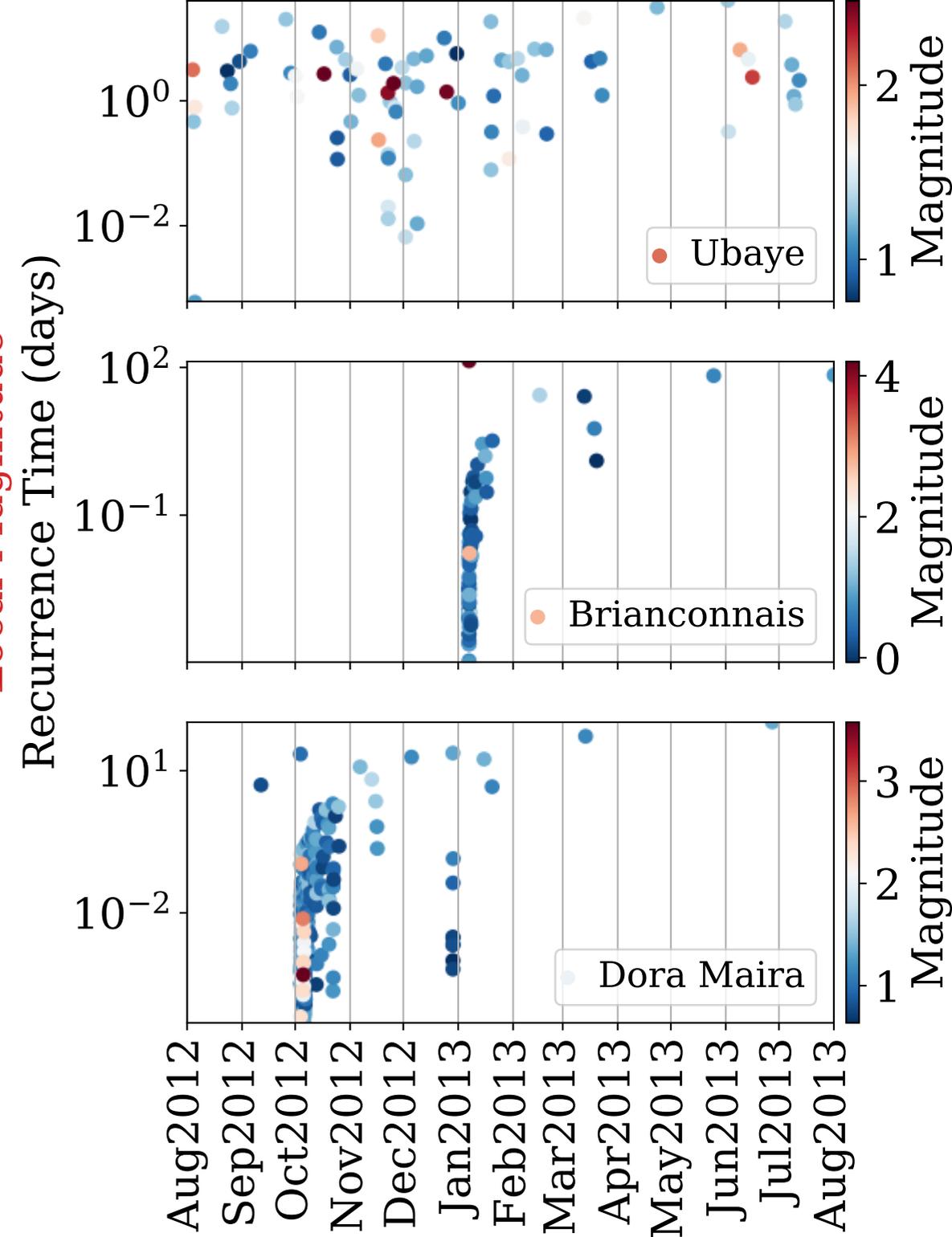
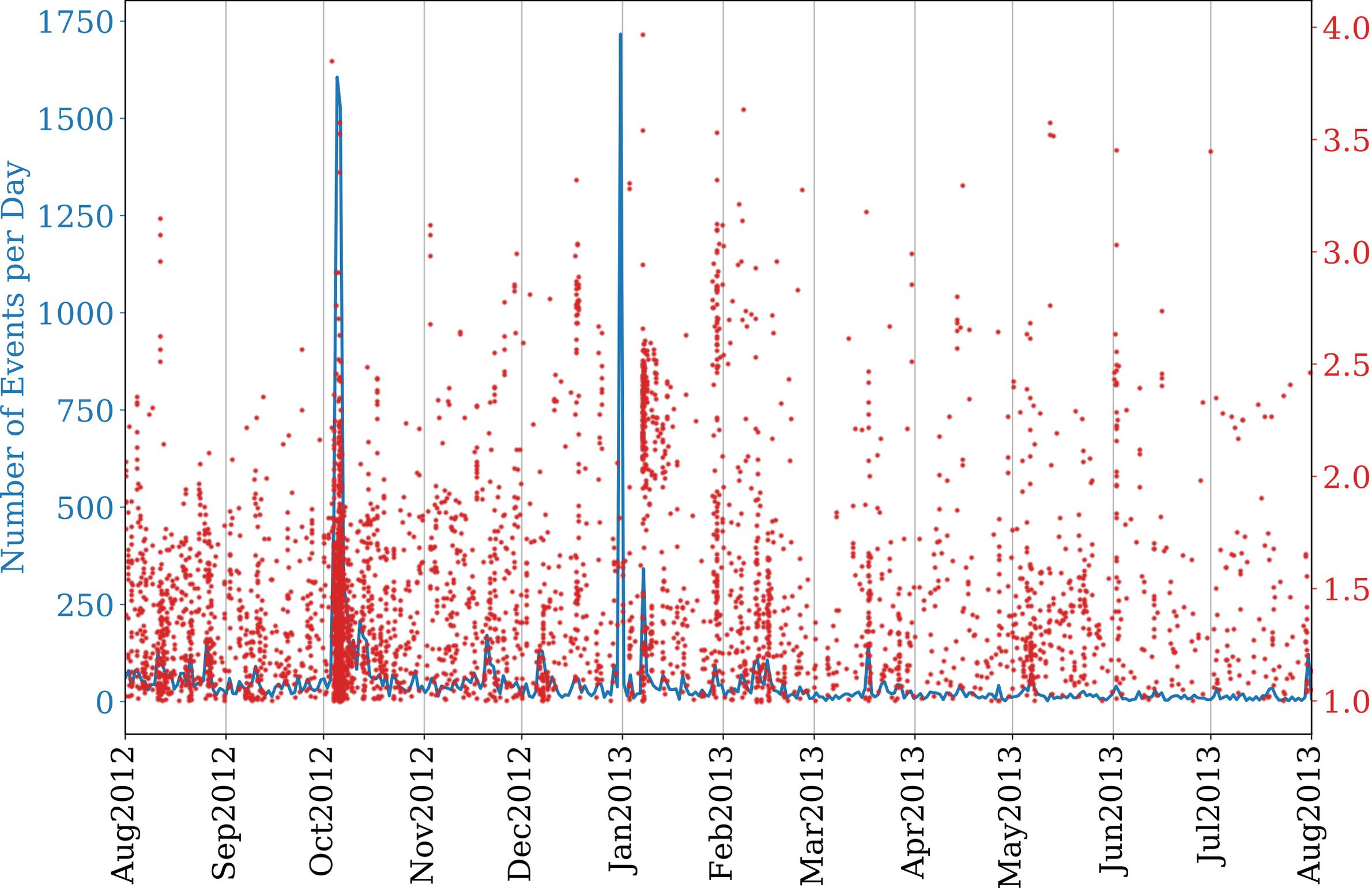


Figure 8.

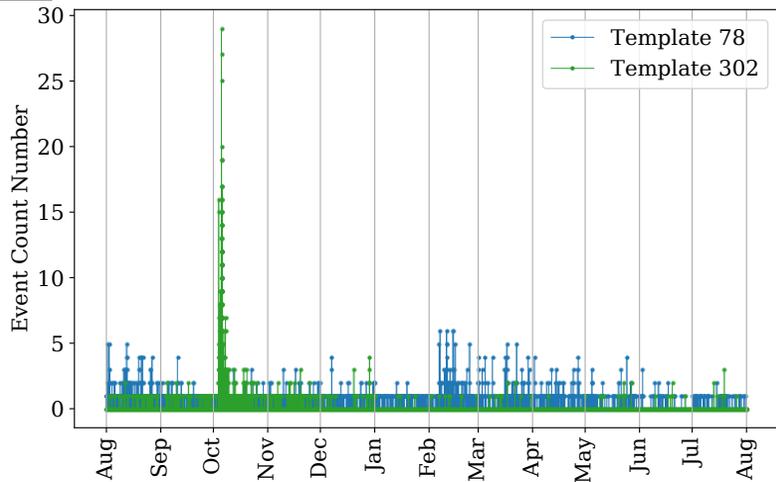
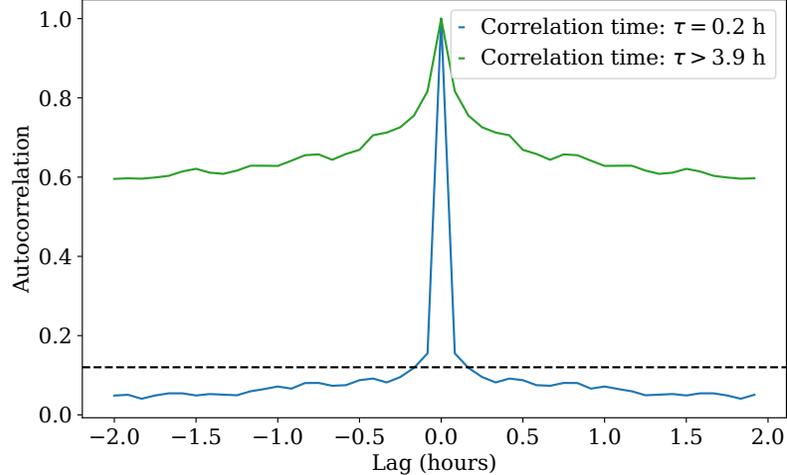
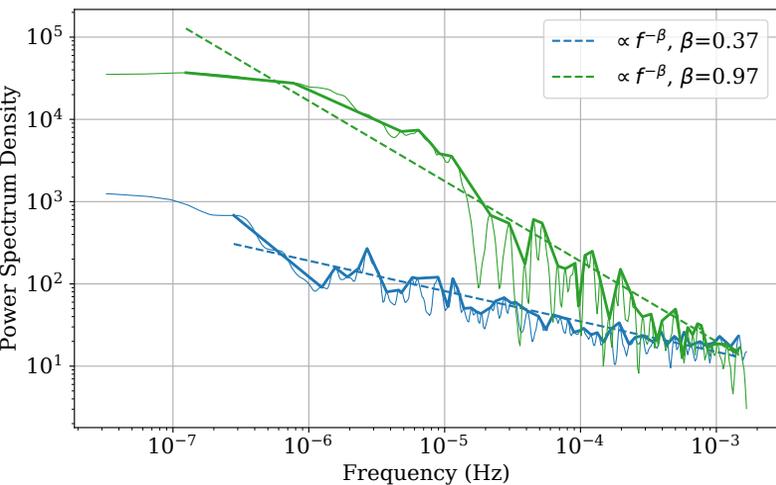
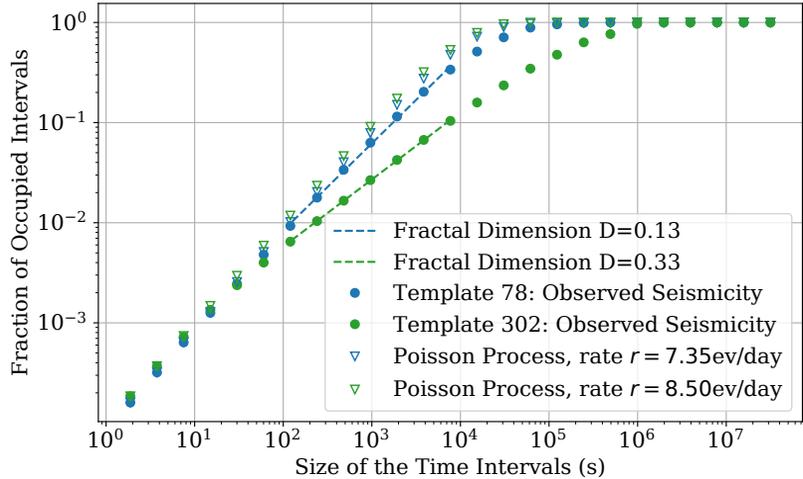
A**B****C****D**

Figure 9.

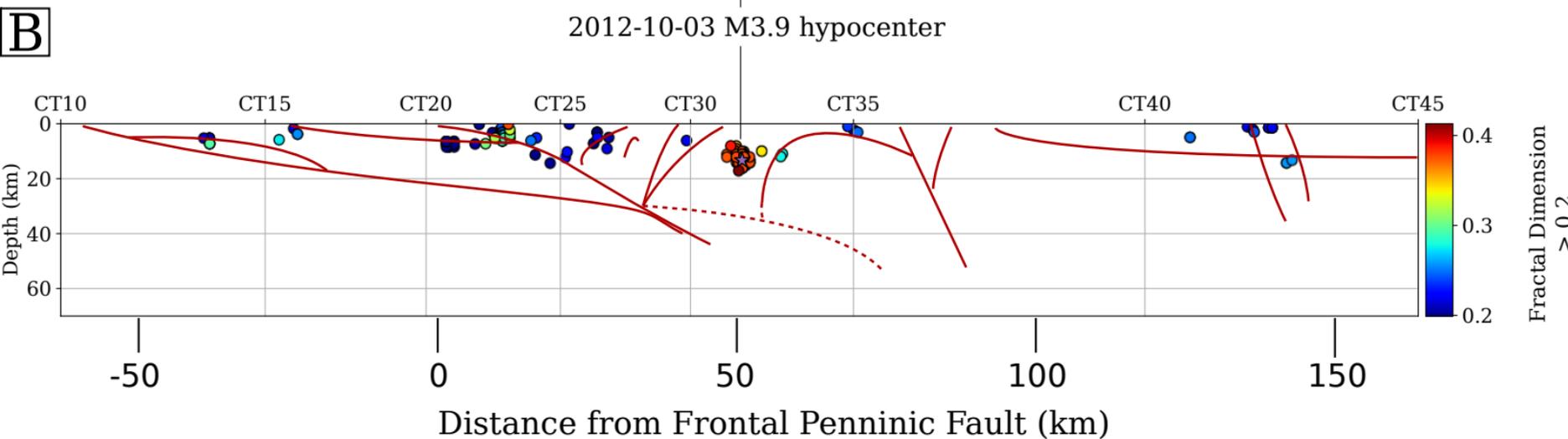
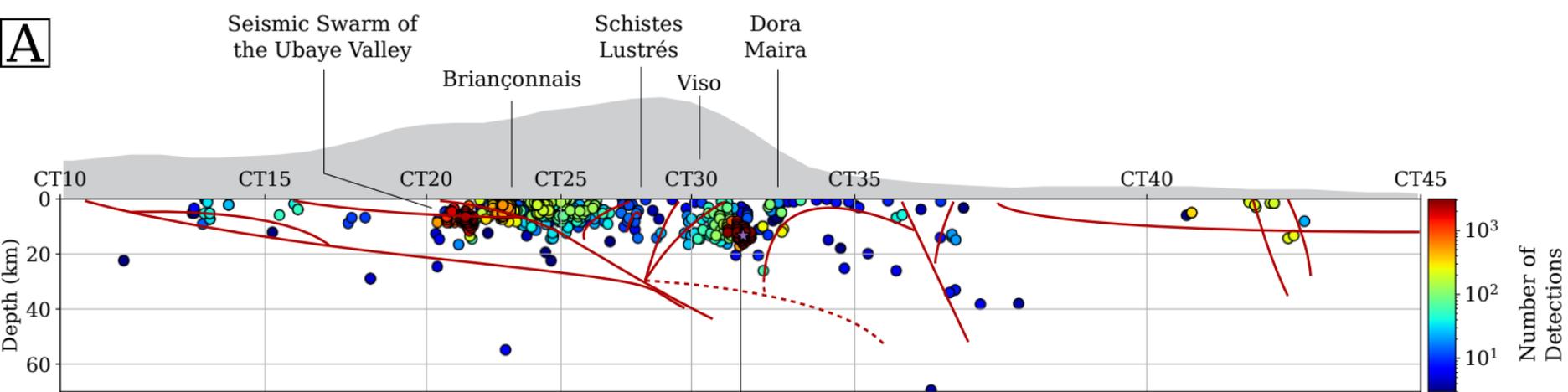


Figure B1.

Template 155: Event 544 (variance reduction = 0.98)

