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Unsupervised Classification of Array Data Based on the L1-Norm

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Abstract—L1-norm criteria have been the subject a flurry of research in signal processing and machine learning over the last decade, especially due to their ability to exploit the sparsity of latent variables and their robustness in the presence of faulty data. Among such criteria, L1-norm principal component analysis (L1-PCA) has drawn considerable attention, resulting in a variety of optimization algorithms and connections with other data processing techniques such as independent component analysis. The present contribution takes a step forward in the characterization of L1-PCA by exploring its linear discrimination capabilities. A variant of L1-PCA consisting of L1-norm maximization subject to an L2-norm constraint is put forward for unsupervised classification. The discrimination properties of the proposed L1-PCA variant are demonstrated through a number of computer experiments.

I. INTRODUCTION

Data analysis techniques based on the L1 norm have become ubiquitous in signal processing and machine learning. Indeed, L1-norm costs arise as natural criteria for exploiting the underlying sparsity in many real-world problems and present increased robustness to faulty data as compared with traditional L2-norm based techniques such as second-order principal component analysis (L2-PCA) [1].

A remarkable L1-norm approach is L1-PCA [2], [3]. While behaving like L2-PCA in many situations, L1-PCA offers enhanced immunity against outliers as well as an interesting flexibility in analyzing data patterns. These advantages come at the expense of numerical complexity, as the absolute value is difficult to treat mathematically, making the optimization of L1-norm criteria more computationally demanding than that of traditional L2-norm criteria. To alleviate this difficulty, a number of optimal and suboptimal algorithms have been proposed for computing L1-PCA [2], [3], [4].

Although increasing complexity, the absolute value nonlinearity confers great potential to this technique. The nonlinearity introduces higher-order statistics in the cost function, implying, in particular, that L1-PCA can perform independent component analysis (ICA) in a robust manner under rather general assumptions [5]. This finding opens interesting possibilities for the use of L1-PCA in application scenarios for which it was not originally conceived, as the technique yields meaningful results under different underlying data models.

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Motivated by the flexibility of this technique, the present work investigates for the first time L1-PCA in a data classification framework. Traditional L2-PCA has been used in unsupervised classification tasks, a fundamental data analysis problem. Unsupervised classification aims at allocating an observed feature vector to one of different classes according to a decision rule based on some proximity criterion. In the fundamental two-class (or binary) case, projecting the observed samples in the direction that maximizes the projected data variance, i.e., applying L2-PCA, can perform unsupervised classification when the data form two well-separated clusters. As a prime example, the *Eigenfaces* method for face recognition is largely based on this property [6], [7]. Inspired by this result, we move forward in the characterization of L1-PCA by analyzing its discrimination capabilities. An L1-PCA variant is proposed that is able to carry out unsupervised classification, much like its L2-norm counterpart but with increased discrimination properties and robustness to outliers. These results are justified using intuitive arguments and supported by several numerical experiments.

II. PROBLEM FORMULATION

This section summarizes the main assumptions and mathematical notations used throughout the paper. Then L2-PCA and L1-PCA are recalled, as well as the linear discrimination properties of the former. These inspire the search for analogous properties in L1-norm criteria.

A. Assumptions and Notations

The data analysis techniques considered in this paper assume the observation of N samples $\mathbf{x}_1, \mathbf{x}_2, \ldots, \mathbf{x}_N$ drawn from a *d*-dimensional random variable $\mathbf{x} \in \mathbb{R}^d$. For simplicity, it is assumed that $\{\mathbf{x}_n\}_{n=1}^N$ has zero sample mean. In the sequel, we use the standard definitions of the L1-norm and the L2-norm, respectively, of an N-dimensional vector $\mathbf{y} \in \mathbb{R}^N$:

$$\|\mathbf{y}\|_{1} = \sum_{n=1}^{N} |y_{n}| \qquad \|\mathbf{y}\|_{2} = \sqrt{\left(\sum_{n=1}^{N} y_{n}^{2}\right)}.$$
 (1)

Also, symbol $(\cdot)^{\top}$ denotes the matrix transpose operator and diag (a_1, a_2, \ldots, a_K) the $K \times K$ diagonal matrix with diagonal entries $\{a_k\}_{k=1}^K$.

B. L2-PCA

We first recall the traditional L2-PCA. Its goal is to find a direction $\mathbf{a} \in \mathbb{R}^d$, $\|\mathbf{a}\|_2 = 1$, such that the projected samples

$$y_n = \mathbf{a}^\top \mathbf{x}_n \tag{2}$$

have maximum dispersion as measured by variance. This is also the linear projection best fitting the data in the leastsquares sense [1]. By defining the vector of projected samples as $\mathbf{y} = [y_1, y_2, \dots, y_N]^{\top}$, the problem takes in practice the following form:

$$\max_{\|\mathbf{a}\|_{2}=1} \|\mathbf{y}\|_{2} = \max_{\|\mathbf{a}\|_{2}=1} \sum_{n=1}^{N} \left(\mathbf{a}^{\top} \mathbf{x}_{n}\right)^{2}$$
(3)

where we are using the sample variance instead of the population variance to introduce the optimization criterion. By defining the $d \times N$ sample matrix as $\mathbf{X} = [\mathbf{x}_1, \dots, \mathbf{x}_N]$, we can write $\mathbf{y} = \mathbf{X}^\top \mathbf{a}$, and the above criterion accepts the equivalent matrix formulation:

$$\max_{\|\mathbf{a}\|_2=1} \|\mathbf{X}^{\top}\mathbf{a}\|_2$$

Simple algebra shows that this optimization problem is solved by the eigenvector associated with the largest eigenvalue of the data sample covariance matrix

$$\mathbf{C}_{\mathbf{x}} = \mathbf{X}\mathbf{X}^{\top}.$$
 (4)

C. L1-PCA

A major problem with the above technique is that the square in (3) tends to overemphasize large outliers and L2-PCA inherits this sensitivity to faulty data. To overcome this drawback, the square function can be replaced by the absolute value, as proposed in [2], thus yielding the following criterion:

$$\max_{\|\mathbf{a}\|_{2}=1} \|\mathbf{y}\|_{1} = \max_{\|\mathbf{a}\|_{2}=1} \|\mathbf{X}^{\top}\mathbf{a}\|_{1}.$$
 (5)

PCA based on criterion (5) is usually referred to as *L1-norm* based PCA or, simply, *L1-PCA*. The absolute value in eqn. (5) endows L1-PCA with an increased robustness to outliers as compared to L2-PCA.

The main difficulty with L1-PCA lies in the optimization of its cost function, which requires the development of specific algorithms. A cost-effective though suboptimal method is proposed by Kwak [2], who proves that the following fixedpoint iteration:

1)
$$\mathbf{y} = \mathbf{X}^{\top} \mathbf{a}_i$$

2) $\mathbf{a}_{i+1} = \frac{\mathbf{X} \operatorname{sign}(\mathbf{y})}{\|\mathbf{X} \operatorname{sign}(\mathbf{y})\|_2}$

where \mathbf{a}_i represents the update of \mathbf{a} at the *i*th iteration, monotonically increases the L1-norm criterion $\|\mathbf{y}\|_1$ after each step, and so the algorithm converges at least to a local maximum. The most notable feature of this simple method is that no parameters need to be tuned. Alternatively, one can use the polynomial time approaches in [8], [9], [10]. L1-PCA algorithms with guaranteed convergence to a global maximum have also been proposed, although they come at the expense of increased computational complexity [3], [11]. In addition, a simplified and faster, yet suboptimal, version of [3] can be found in [4].

D. Unsupervised Classification via L2-PCA

We now recall that L2-PCA can perform unsupervised classification or *clustering* in the two-class scenario when the clusters are well separated [6]. To see this, let us assume a binary data classification framework where each sample \mathbf{x}_n belongs to one of two possible classes C_k , k = 1, 2, characterized by sample means (or centroids) and covariance matrices:

$$\boldsymbol{\mu}_k = \frac{1}{N_k} \sum_{\mathbf{x}_n \in \mathcal{C}_k} \mathbf{x}_n \qquad \mathbf{C}_k = \frac{1}{N_k} \sum_{\mathbf{x}_n \in \mathcal{C}_k} (\mathbf{x}_n - \boldsymbol{\mu}_k) (\mathbf{x}_n - \boldsymbol{\mu}_k)^\top.$$

L2-PCA searches for the dominant eigenvector of the overall data covariance matrix C_x (4). Under our working assumptions, this matrix can be expressed up to an irrelevant scale factor as:

$$\mathbf{C}_{\mathbf{x}} = \sum_{k=1}^{2} \pi_k \mathbf{C}_k + \pi_1 \pi_2 \Delta \boldsymbol{\mu} \Delta \boldsymbol{\mu}^{\top}$$

where $\pi_k = N_k/N$ and $\Delta \mu = (\mu_2 - \mu_1)$ is the line joining the class means. If the data dispersion in each class is negligible compared to the distance between the class means, i.e., $\|\mathbf{C}_k\|_{\text{Fro}} \ll \|\Delta \mu\|_2$, where $\|\cdot\|_{\text{Fro}}$ denotes the Frobenius norm, then we can approximate

$$\mathbf{C}_{\mathbf{x}} \approx \pi_1 \pi_2 \Delta \boldsymbol{\mu} \Delta \boldsymbol{\mu}^{\top}.$$

Consequently, L2-PCA defines a projection in the direction $\Delta \mu$ joining the class centroids.

To illustrate why this direction may be useful for classification, Figure 1 represents the scatter plot of a zero-mean synthetic data set with two clusters ('green' and 'red'), as well as the data projections (2) along unit vector **a** in the direction of the line joining the cluster means. We can see that the projection yields mostly negative values for the green class and mostly positive for the red class, thus allowing cluster discrimination based on these values.

III. L1-NORM BASED DISCRIMINATION

Pursuing these intuitive arguments, we can define an L1-PCA based criterion for unsupervised clustering as follows.

A. Separating the Projected Class Centroids

In the specific case illustrated in Figure 1, we observe that the sign of the projection gives us clues about the class label: a data vector is most likely 'red' if its projection is positive, and 'green' otherwise. The better the classes are separated after the projection, the more accurate the estimation but, in any case, the sign allows us to eliminate any reference to the 'colors' or categories of the data points. This is therefore an unsupervised approach. Now, the L1 norm recalled in eqn. (1) is, by definition, the sum of the positively signed projections minus the sum of the negatively signed ones:

$$\|\mathbf{y}\|_1 = \sum_{n=1}^N |y_n| = \sum_{y_n>0} y_n - \sum_{y_n<0} y_n.$$



Fig. 1: Scatter plot of samples from two classes ('green' and 'red'). The small vertical lines on the horizontal axis represent the sample projections onto the dashed line joining the class means Also shown are the histograms and kernel estimates of the densities of the projected data.

In a situation such as that depicted in Figure 1, we can think of this difference as an indicator of the distance between the projected clusters, so that the more separated the clusters, the larger the L1-norm. Therefore, maximizing criterion (5) appears as a promising option for classification. There is an additional interpretation: in many cases, the direction of maximum variance produces well separated clusters; this is why traditional PCA has discriminative properties as recalled in the previous section. Here, we simply reformulate this idea by substituting the variance with the L1-norm for measuring the spread of the data. It is implicitly assumed that the clusters are sufficiently distant in the original *d*-dimensional space, but this assumption is reasonable in many classification problems.

In practice, however, it may not be sufficient to just search for a vector **a** such that the corresponding projections y in (2) have a large variance. It is for this reason that L2-PCA fails when the class dispersion is significant relative to the centroid distance, even if the clusters are linearly separable [7]. To prevent the overlapping between the projected clusters, the spread of the projections around their respective centroids should also be minimized, much like in classical linear discriminant analysis (LDA) [12].

B. Reducing the Projected Data Dispersion

The scatter can be measured by $\sigma^2_{\rm pos}$ and $\sigma^2_{\rm neg},$ defined as

$$\sigma_{\text{pos}}^{2} = \frac{1}{N_{\text{pos}}} \sum_{y_{n}>0} (y_{n} - m_{\text{pos}})^{2} \quad m_{\text{pos}} = \frac{1}{N_{\text{pos}}} \sum_{y_{n}>0} y_{n}$$
$$\sigma_{\text{neg}}^{2} = \frac{1}{N_{\text{neg}}} \sum_{y_{n}<0} (y_{n} - m_{\text{neg}})^{2} \quad m_{\text{neg}} = \frac{1}{N_{\text{neg}}} \sum_{y_{n}<0} y_{n}$$

where $N_{\rm pos}$ and $N_{\rm neg}$ represent the number of positive and negative projections, respectively, with $N_{\rm pos} + N_{\rm neg} = N$. The weighted sum $N_{\rm pos}\sigma_{\rm pos}^2 + N_{\rm neg}\sigma_{\rm neg}^2$ can be interpreted as an estimate of the total spread of the projected data. Some algebra shows that

$$\sigma_{\rm pos}^2 = \frac{1}{N_{\rm pos}} \sum_{y_n > 0} y_n^2 - m_{\rm pos}^2$$
$$\sigma_{\rm neg}^2 = \frac{1}{N_{\rm neg}} \sum_{y_n < 0} y_n^2 - m_{\rm neg}^2.$$

For zero-mean data, it also holds that

$$\sum_{y_n>0} y_n = -\sum_{y_n<0} y_n \Rightarrow N_{\text{neg}} m_{\text{neg}} = -N_{\text{pos}} m_{\text{pos}}$$

and

$$\|\mathbf{y}\|_1 = \sum_{n=1}^N |y_n| = \sum_{y_n > 0} y_n - \sum_{y_n < 0} y_n = 2N_{\text{pos}} m_{\text{pos}}.$$

Putting it all together, it can be shown that

$$\begin{split} N_{\rm pos}\sigma_{\rm pos}^2 + N_{\rm neg}\sigma_{\rm neg}^2 &= \|\mathbf{y}\|_2^2 - \mu \|\mathbf{y}\|_1^2 \end{split}$$
 where $\mu = \frac{N}{4N_{\rm pos}N_{\rm neg}} > 0.$

C. L1-norm Criterion for Unsupervised Classification

As a result of the above developments, we can state that:

- To maximize the distance between the projected data centroids, one should maximize $\|\mathbf{y}\|_1$.
- To minimize the dispersion around the projected data centroids, one should maximize $\|\mathbf{y}\|_1$ while minimizing $\|\mathbf{y}\|_2$ or at least keeping it constant.

Accordingly, the following Rayleigh quotient arises as a sensible criterion for unsupervised classification based on the L1 norm:

$$\max_{\mathbf{a}} J_1(\mathbf{a}) = \frac{\|\mathbf{y}\|_1}{\sqrt{\|\mathbf{y}\|_2}}.$$
 (6)

The purpose of taking the square root of the L2 norm in the denominator of (6) is to ensure that the objective function is dimensionless and invariant to the scale of \mathbf{a} , i.e., $J_1(\lambda \mathbf{a}) = J_1(\mathbf{a}), \lambda \in \mathbb{R} \setminus \{0\}$. This criterion is experimentally assessed next.

IV. EXPERIMENTAL RESULTS

This section evaluates the performance of the proposed L1norm criterion through several computer experiments.

A. Discriminative Properties

The first experiment randomly draws N = 500 samples from a mixture of two equiprobable ($\pi_1 = \pi_2 = 0.5$) twodimensional Gaussian distributions with means $\boldsymbol{\mu}_1 = -[2,0]^{\top}$ and $\boldsymbol{\mu}_2 = [2,0]^{\top}$ and covariance matrices $\mathbf{C}_1 = \mathbf{C}_2 =$ diag(1,3). Figure 2 depicts the proposed criterion $J_1(\mathbf{a})$, defined in eqn. (6), as a function of the angle θ between the projection vector \mathbf{a} and the x-axis (blue line). The curve attains its global maximum at approximately $\theta = 0$ degrees, i.e., the optimal direction for this dataset is parallel to the horizontal axis. For completeness, Figure 2 also represents the



Fig. 2: Proposed criterion $J_1(\theta)$ defined in eqn. (6) (blue line) for $y = \mathbf{a}^\top \mathbf{x}$ with $\mathbf{a} = [\cos(\theta), \sin(\theta)]^\top$, $\theta \in [-\pi, \pi]$. For comparison, the L2-PCA objective function (3) is also depicted (black line). Both curves are normalized with respect to their maximum values.

variance of the projected data (black line), which corresponds to criterion (3) maximized by L2-PCA.

The scatter plot of the observations is given in Figure 3, where data of class 1 and class 2 are represented by circles and crosses, respectively. The blue line points in the direction maximizing $J_1(\mathbf{a})$. As reported in the previous paragraph, observe that this line is almost horizontal. Compare with the yellow line, which is the direction determined by classical LDA [12]: there is no significant difference between the projections calculated by the two methods, as both lines practically overlap. Finally, we point out that the black dashed line marked PCA in Figure 3 points in the direction of maximum variance of the data. Even though L2-PCA seems to yield an acceptable result, i.e., the projected data form clusters, this technique may have poor performance when the data dispersion is large compared to the distance between the class means. This is clearly illustrated by Figure 4, which shows the result obtained by repeating the experiment with class covariance matrices $C_1 = C_2 = diag(1, 10)$, so that the direction of maximum variance is now along the x_2 -axis. Obviously, projecting along this axis does not produce well separated groups. By contrast, the proposed criterion is essentially unaffected by the change in experimental conditions.

B. Robustness Against Outliers

Outliers are samples generated by a different mechanism, or model, than the nominal observations. In this experiment, we replace 33% of the data in Figure 3 by outliers generated by raising a Gaussian random variable to the third power. The mean of the outliers in each class is set to that of the data they are replacing. Figure 5 represents the scatter plot of the corrupted dataset. The blue line again points in the direction maximizing the proposed criterion J_1 , while the black line represents the direction of maximum variance of the data. The experiment shows that, as expected, the proposed L1-norm based method is more robust against outliers than



Fig. 3: Performance of the proposed algorithm when applied to classes with almost spherical data point clouds. Results obtained by using L2-PCA and LDA are also shown for comparison.



Fig. 4: Performance of the proposed algorithm when applied to classes with non-spherical data point clouds. Results obtained by using L2-PCA and LDA are also shown for comparison.

L2-PCA. Figure 6 plots the estimated probability density function, computed by a kernel density estimator method, of the projection of the data onto the direction maximizing J_1 (blue line of Figure 5). The distribution is clearly bimodal, showing that the sample arises from two different classes.

As a further experiment, we fit a Gaussian mixture model (GMM) to the corrupted data. GMMs are based on the assumption that all the observations are drawn from a mixture of normal distributions, and are very popular for solving unsupervised classification problems [13]. Roughly speaking, GMMs generalize the standard k-means clustering algorithm by incorporating information about the covariance structure of the data classes. In this experiment, the parameters of the model are found by using the Expectation-Maximization (EM) algorithm [13]. Figure 7 plots the data over the es-



Fig. 5: Performance of the proposed algorithm and L2-PCA when applied to data corrupted by outliers.



Fig. 6: Histogram of the projection of the corrupted data onto the optimal L1-norm based direction, shown by the blue line of Figure 5.

timated Gaussian mixture model contours. Observe that the EM algorithm models the dataset as a mixture of two almostoverlapped Gaussians with *a priori* probabilities $\pi_1 = 0.94$ and $\pi_2 = 0.06$. After learning the GMM, a data point is classified according to the probability of having been generated by each of the Gaussians. Figure 7 also plots the predicted labels of the data points: data from class 1 are shown as green dots, while data from class 2 are shown as red crosses. The resulting classification is clearly erroneous and, in view of Figure 6, it can easily be outperformed by the proposed L1-PCA based approach.

V. CONCLUSIONS

The present contribution has framed L1-PCA in the context of data classification. An L1-PCA variant has been derived that can perform unsupervised linear discrimination in the twoclass setting. The variant consists of the L1-norm of the vector of projected data samples normalized by the square root of its L2 norm, which can be cast as a Rayleigh quotient. Although the criterion has been derived on intuitive grounds, a number of



Fig. 7: Contour plot of the estimated Gaussian mixture model of the corrupted data. Tentative classification of the data points as suggested by this model is also shown, yielding highly inaccurate results.

computer experiments have confirmed its good discrimination capabilities and enhanced robustness to outliers as compared to classical L2-PCA. Further work should include a thorough theoretical study and a comparative performance analysis on real data.

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