MASSICCC: A SaaS Platform for Clustering and Co-Clustering of Mixed Data

https://massiccc.lille.inria.fr/

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21-22 June 2018, Workshop MixStatSeq, Paris "Mixture models: Theory and applications"







Outline

- 1 Introduction
- 2 Model-based clustering
- 3 Mixmod in MASSICCO
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCO
- 6 Conclusion

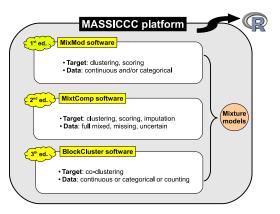
MASSICCC?

massiccc.lille.inria.fr



SaaS: Software as a Service

MASSICCC??



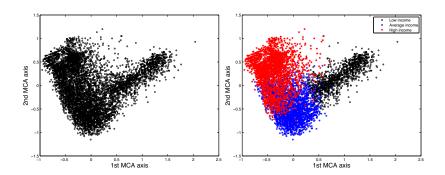
A high quality and easy to use web platform where are transfered mature research clustering (and more) software towards (non academic) professionals

Here is the computer you need!

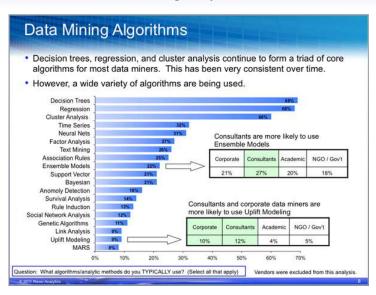


Clustering?

Detect hidden structures in data sets

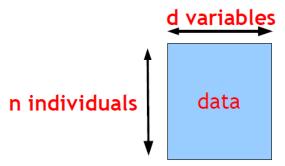


Clustering everywhere¹

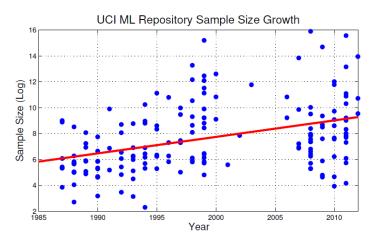


¹Rexer Analytics's Annual Data Miner Survey is the largest survey of data mining, data science, and analytics professionals in the industry (survey of 2011)

Data sets structure

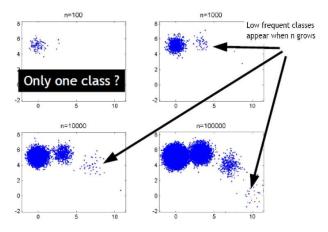


Large data sets²

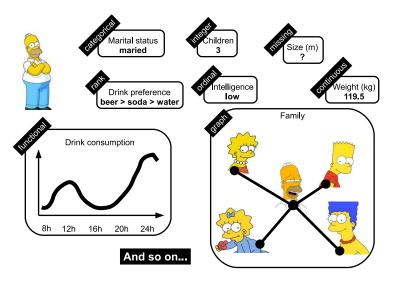


²S. Alelyani, J. Tang and H. Liu (2013). Feature Selection for Clustering: A Review. *Data Clustering: Algorithms and Applications*, **29**

An opportunity for detecting weak signal



Todays features: full mixed/missing



Notations

- Data: n individuals: $\mathbf{x} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\} = \{\mathbf{x}^O, \mathbf{x}^M\}$ in a space \mathcal{X} of dimension d
 - Observed individuals x^O
 Missing individuals x^M
- \blacksquare Aim: estimation of the partition **z** and the number of clusters KPartition in K clusters G_1, \ldots, G_K : $\mathbf{z} = (\mathbf{z}_1, \ldots, \mathbf{z}_n), \ \mathbf{z}_i = (z_{i1}, \ldots, z_{iK})'$

$$\mathbf{x}_i \in G_k \quad \Leftrightarrow \quad \mathbf{z}_{ih} = \mathbb{I}_{\{h=k\}}$$

Mixed, missing, uncertain								
Individuals x				Partition z			\Leftrightarrow	Group
?	0.5	red	5	?	?	?	\Leftrightarrow	???
0.3	0.1	green	3	?	?	?	\Leftrightarrow	???
0.3	0.6	{red,green}	3	?	?	?	\Leftrightarrow	???
0.9	[0.25 0.45]	red	?	?	?	?	\Leftrightarrow	???
+	.	\downarrow	\downarrow					
continuous	continuous	categorical	integer					
				•				

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Parametric mixture model

■ Parametric assumption:

$$p_k(\mathbf{x}_1) = p(\mathbf{x}_1; \boldsymbol{\alpha}_k)$$

thus

$$p(\mathbf{x}_1) = p(\mathbf{x}_1; \boldsymbol{\theta}) = \sum_{k=1}^K \pi_k p(\mathbf{x}_1; \boldsymbol{\alpha}_k)$$

■ Mixture parameter:

$$heta=(oldsymbol{\pi},oldsymbol{lpha})$$
 with $oldsymbol{lpha}=(oldsymbol{lpha}_1,\ldots,oldsymbol{lpha}_K)$

lacktriangle Model: it includes both the family $p(\cdot; \alpha_k)$ and the number of groups K

$$\mathbf{m} = \{ \mathsf{p}(\mathbf{x}_1; \boldsymbol{\theta}) : \boldsymbol{\theta} \in \Theta \}$$

The number of free continuous parameters is given by

$$\nu = \dim(\Theta)$$

Clustering becomes a well-posed problem...

The clustering process in mixtures

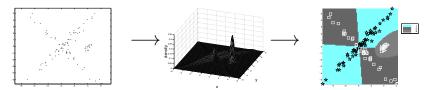
- **I** Estimation of θ by $\hat{\theta}$
- f Z Estimation of the conditional probability that $f x_i \in G_k$

$$t_{ik}(\hat{\boldsymbol{\theta}}) = p(Z_{ik} = 1 | \mathbf{X}_i = \mathbf{x}_i; \hat{\boldsymbol{\theta}}) = \frac{\hat{\pi}_k p(\mathbf{x}_i; \hat{\boldsymbol{\alpha}}_k)}{p(\mathbf{x}_i; \hat{\boldsymbol{\theta}})}$$

SESTIMATION of **z**_i by maximum a posteriori (MAP)

$$\hat{z}_{ik} = \mathbb{I}_{\{k = \operatorname{arg} \max_{h=1,...,K} t_{ih}(\hat{\theta})\}}$$

4 Model selection: BIC, ICL, ...

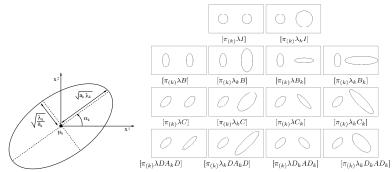


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Only continuous features: 14 models on Σ_k





Only categorical variables: latent class model

lacktriangle Categorical variables: d variables with m_j modalities each, $\mathbf{x}_i^j \in \{0,1\}^{m_j}$ and

$$\mathbf{x}_i^{jh} = 1 \quad \Leftrightarrow \quad \text{variable } j \text{ of } \mathbf{x}_i \text{ takes level } h$$

■ Conditional independence:

$$p(\mathbf{x}_i; \boldsymbol{\alpha}_k) = \prod_{j=1}^d \prod_{h=1}^{m_j} (\alpha_k^{jh})^{x_i^{jh}}$$

and

$$\alpha_k^{jh} = p(\mathbf{x}_i^{jh} = 1|z_{ik} = 1)$$

with
$$\alpha_k = (\alpha_k^{jh}; j = 1, ..., d; h = 1, ..., m_j)$$

Mixing continuous and categorical data: full local independence

Combine continuous and categorical data

$$extbf{ extit{x}}_1 = (extbf{ extit{x}}_1^{cont}, extbf{ extit{x}}_1^{cat})$$

The proposed solution is to mixed both types by inter-type conditional independence

$$p(\mathbf{x}_1; \boldsymbol{\alpha}_k) = p(\mathbf{x}_1^{cont}; \boldsymbol{\alpha}_k^{cont}) \times p(\mathbf{x}_1^{cat}; \boldsymbol{\alpha}_k^{cat})$$

In addition, for symmetry between types, intra-type conditional independence

Only need to define the univariate pdf for each variable type!

- Continuous: Gaussian
- Categorical: multinomial

Estimation of heta by *complete*-likelihood

Maximize the complete-likelihood over (θ, z)

$$\ell_c(\boldsymbol{\theta}; \mathbf{x}, \mathbf{z}) = \sum_{i=1}^n \sum_{k=1}^K z_{ik} \ln \left\{ \pi_k \mathbf{p}(\mathbf{x}_i; \boldsymbol{\alpha}_k) \right\}$$

■ Equivalent to traditional methods

Metric	M = I	M free	\mathbf{M}_k free		
Gaussian model	$[\pi \lambda I]$	$[\pi \lambda C]$	$[\pi \lambda_k C_k]$		

- lacksquare Bias of $\hat{ heta}$: heavy if poor separated clusters
- Associated optimization algorithm: CEM (see later)
- CEM with $[\pi \lambda I]$ is strictly equivalent to K-means
- CEM is simple et fast (convergence with few iterations)

Maximize the observe-likelihood on θ

$$\ell(\boldsymbol{\theta}; \mathbf{x}) = \sum_{i=1}^{n} \ln p(\mathbf{x}_i; \boldsymbol{\theta})$$

- lacktriangle Convergence of $\hat{\theta}$, asymptotic efficiency, asymptotically unbiased
- General algorithm for missing data: EM
- EM is simple but slower than CEM
- Interpretation: it is a kind of fuzzy clustering

Principle of EM and CEM

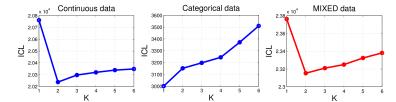
- Initialization: θ^0
- Iteration n^oq:
 - Step E: estimate probabilities $\mathbf{t}^q = \{t_{ik}(\boldsymbol{\theta}^q)\}$
 - Step C: classify by setting $\mathbf{t}^q = \mathsf{MAP}(\{t_{ik}(\boldsymbol{\theta}^q)\})$
 - Step M: maximize $\theta^{q+1} = \operatorname{arg\,max}_{\theta} \ell_c(\theta; \mathbf{x}, \mathbf{t}^q)$
- Stopping rule: iteration number or criterion stability

Properties

- ⊕: simplicity, monotony, low memory requirement
- \blacksquare \ominus : local maxima (depends on θ^0), linear convergence (EM)

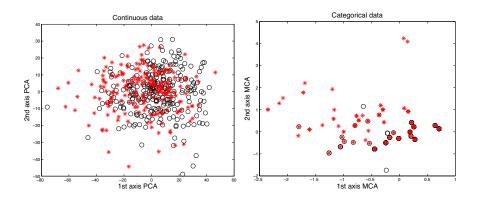
Prostate cancer data (without mixing data)

- Individuals: n = 475 patients with prostatic cancer grouped on clinical criteria into two Stages 3 and 4 of the disease
- \blacksquare Variables: d=12 pre-trial variates were measured on each patient, composed by eight continuous variables (age, weight, systolic blood pressure, diastolic blood pressure, serum haemoglobin, size of primary tumour, index of tumour stage and histolic grade, serum prostatic acid phosphatase) and four categorical variables with various numbers of levels (performance rating, cardiovascular disease history, electrocardiogram code, bone metastases)
- Model: cond. indep. $p(\mathbf{x}_1; \alpha_k) = p(\mathbf{x}_1; \alpha_k^{cont}) \cdot p(\mathbf{x}_1; \alpha_k^{cat})$

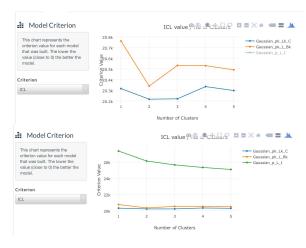


Prostate cancer data (without missing data)

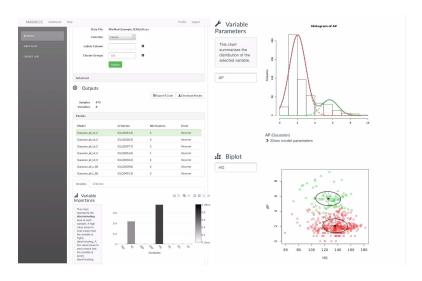
Variables	Continuous		Categ	gorical	Mixed		
Error (%)	9.46		47.16		8.63		
True \ estimated group	1	2	1	2	1	2	
Stage 3	247	26	142	131	252	21	
Stage 4	19	183	120	82	20	182	



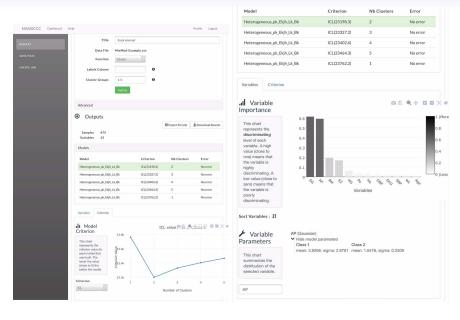
Continuous data



Continuous data



Mixed data



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Full mixed data: conditional independence everywhere³

The aim is to combine continuous, categorical, integer data, ordinal, ranking and functional data

$$\textit{\textbf{x}}_1 = (\textit{\textbf{x}}_1^{\textit{cont}}, \textit{\textbf{x}}_1^{\textit{cat}}, \textit{\textbf{x}}_1^{\textit{int}}, \ldots)$$

The proposed solution is to mixed all types by inter-type conditional independence

$$p(\mathbf{x}_1; \boldsymbol{\alpha}_k) = p(\mathbf{x}_1^{cont}; \boldsymbol{\alpha}_k^{cont}) \times p(\mathbf{x}_1^{cat}; \boldsymbol{\alpha}_k^{cat}) \times p(\mathbf{x}_1^{int}; \boldsymbol{\alpha}_k^{int}) \times \dots$$

In addition, for symmetry between types, intra-type conditional independence

Only need to define the univariate pdf for each variable type!

- Continuous: Gaussian
- Categorical: multinomial
- Integer: Poisson
-

³MixtComp software on the MASSICCC platform: https://massiccc.lille.inria.fr/

Missing data: MAR assumption and estimation

Assumption on the missingness mecanism

Missing At Randon (MAR): the probability that a variable is missing does not depend on its own value given the observed variables.

Observed log-likelihood...

$$\ell(\boldsymbol{\theta}; \mathbf{x}^{\mathrm{O}}) = \sum_{i=1}^{n} \log \left(\sum_{k=1}^{K} \pi_{k} \mathbf{p}(\mathbf{x}_{i}^{\mathrm{O}}; \boldsymbol{\alpha}_{k}) \right) = \ln \left[\sum_{k=1}^{K} \pi_{k} \underbrace{\int_{\mathbf{x}_{i}^{M}} \mathbf{p}(\mathbf{x}_{i}^{\mathrm{O}}, \mathbf{x}_{i}^{M}; \boldsymbol{\alpha}_{k}) d\mathbf{x}_{i}^{M}}_{\mathsf{MAP}, \mathsf{vertice}} \right]$$

A SEM algorithm to estimate θ by maximizing the observed-data log-likelihood

- Initialisation: $\theta^{(0)}$
- Iteration nb a:
 - **E-step:** compute conditional probabilities $p(x^M, z|x^0; \theta^{(q)})$
 - S-step: draw $(\mathbf{x}^{M(q)}, \mathbf{z}^{(q)})$ from $p(\mathbf{x}^M, \mathbf{z} | \mathbf{x}^0; \boldsymbol{\theta}^{(q)})$
 - M-step: maximize $\theta^{(q+1)} = \arg \max_{\theta} \ln p(\mathbf{x}^O, \mathbf{x}^{M(q)}, \mathbf{z}^{(q)}; \theta)$
- Stopping rule: iteration number

Properties: simpler than EM and interesting properties!

- Avoid possibly difficult E-step in an EM
- Classical M steps
- Avoids local maxima
- The mean of the sequence $(\theta^{(q)})$ approximates $\hat{\theta}$
- The variance of the sequence $(\theta^{(q)})$ gives confidence intervals

⁴MixtComp software on the MASSICCC platform: https://massiccc.lille.inria.fr/

Prostate cancer data (with missing data)⁵

- Individuals: 506 patients with prostatic cancer grouped on clinical criteria into two Stages 3 and 4 of the disease
- Variables: *d* = 12 pre-trial variates were measured on each patient, composed by eight continuous variables (age, weight, systolic blood pressure, diastolic blood pressure, serum haemoglobin, size of primary tumour, index of tumour stage and histolic grade, serum prostatic acid phosphatase) and four categorical variables with various numbers of levels (performance rating, cardiovascular disease history, electrocardiogram code, bone metastases)
- Some missing data: 62 missing values ($\approx 1\%$)

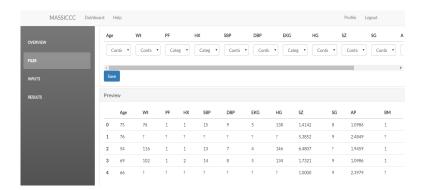
We forget the classes (Stages of the desease) for performing clustering

Questions

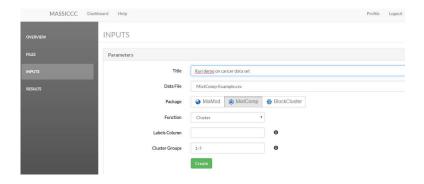
- How many clusters?
- Which partition?

⁵Byar DP, Green SB (1980): Bulletin Cancer, Paris 67:477-488

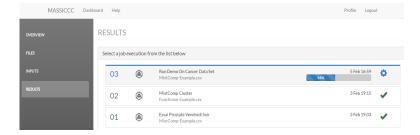
Data upload without preprocessing



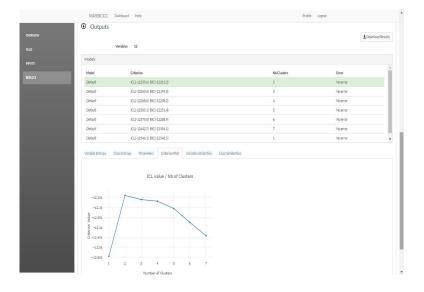
Run clustering analysis



It is running on the (Inria) cloud...



Several quick result overviews... without post-processing

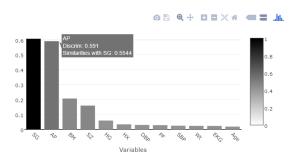


roduction Model-based clustering Mixmod in MASSICCC MixtComp in MASSICCC BlockCluster Conclusion

Variable significance on global partition

Variable Importance

This chart represents the discriminating level of each variable. A high value (close to one) means that the variable is highly discriminating. A low value (close to zero) means that the variable is poorly discriminating. Click on one of the bars to display the distribution of this variable and, to also display the similarities between this variable and all the others. The color of the bars reflects the similarities between all the variables and the selected variable. Read more

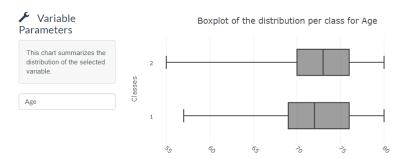


Sort Variables : ↓↑

+ similarity between variables

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Variable "Age" difference between clusters



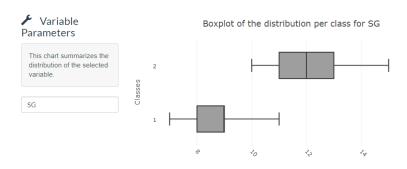
Age (Gaussian)

✔ Hide model parameters

Class 1 Class 2
mean: 71.534, sigma: 6.760 mean: 71.313, sigma: 7.463

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Variable "SG" difference between clusters



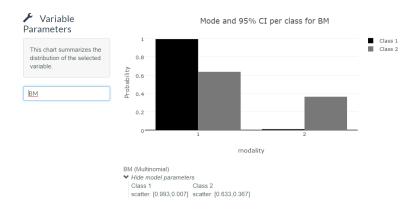
SG (Gaussian)

➤ Hide model parameters

Class 1 Class 2
mean: 8.940, sigma: 1.154 mean: 12.087, sigma: 1.405

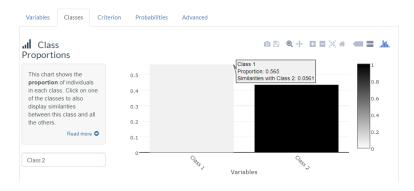
oduction Model-based clustering Mixmod in MASSICCC MixtComp in MASSICCC BlockCluster Conclusio

Variable "BM" difference between clusters



roduction Model-based clustering Mixmod in MASSICCC MixtComp in MASSICCC BlockCluster Conclusion

Individual cluster separation (with the cluster weight)



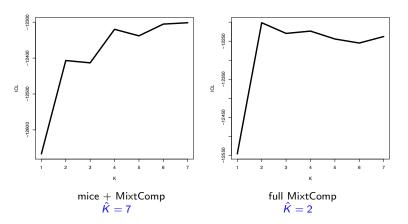
Two strategies in competition

- Strategy "mice⁶ + MixtComp": MixtComp on the dataset completed by mice
- > data.imp=mice(data)
- > data.comp.mice=complete(data.imp)
- Strategy "full MixtComp": MixtComp on the observed (no completed) dataset

Partition quality with K = 2

Strategy	mice + MixtComp	full MixtComp
% misclassified	12.8	8.1

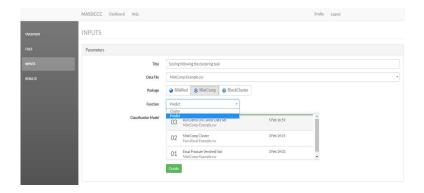
Choosing K with the ICL criterion



... may lose some cluster information when imputation before clustering

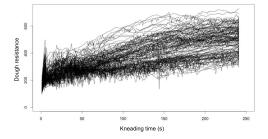
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Scoring cancer data following the clustering task

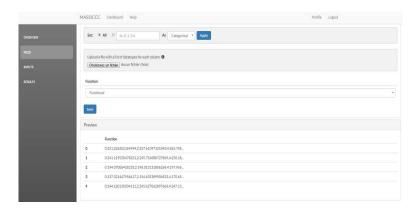


Curve "cookies" data set

The Kneading dataset comes from Danone Vitapole Paris Research Center and concerns the quality of cookies and the relationship with the flour kneading process⁷. There are 115 different flours for which the dough resistance is measured during the kneading process for 480 seconds. One obtains 115 kneading curves observed at 241 equispaced instants of time in the interval [0; 480]. The 115 flours produce cookies of different quality: 50 of them have produced cookies of good quality, 25 produced medium quality and 40 low quality.

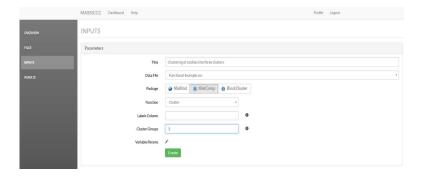


Upload curves data



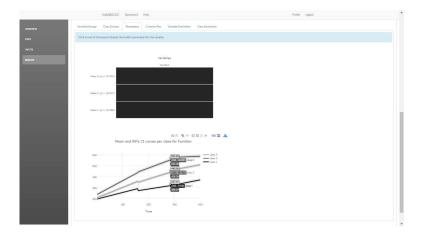
roduction Model-based clustering Mixmod in MASSICCC MixtComp in MASSICCC BlockCluster Conclusion

Run a clustering task with three clusters



troduction Model-based clustering Mixmod in MASSICCC MixtComp in MASSICCC BlockCluster Conclusion

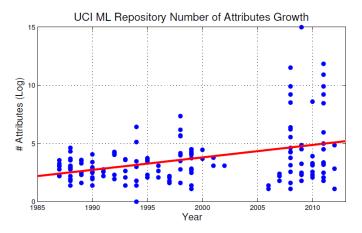
Overview of the three clusters of cookies



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High-dimensional (HD) data⁸



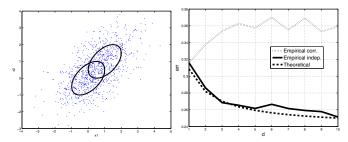
⁸S. Alelyani, J. Tang and H. Liu (2013). Feature Selection for Clustering: A Review. *Data Clustering: Algorithms and Applications*, **29**

Bias/variance in HD: reduce variance, accept bias

A two-component *d*-variate Gaussian mixture with intra-dependency:

$$\pi_1 = \pi_2 = rac{1}{2}, \quad \mathbf{X}_1 | z_{11} = 1 \sim \mathsf{N}_d(\mathbf{0}, \mathbf{\Sigma}), \quad \mathbf{X}_1 | z_{12} = 1 \sim \mathsf{N}_d(\mathbf{1}, \mathbf{\Sigma})$$

- Each variable provides equal and own separation information
- Theoretical error decreases when d grows: $\text{err}_{theo} = \Phi(-\|\mu_2 \mu_1\|_{\Sigma^{-1}}/2)$
- Empirical error rate with the (true) intra-correlated model worse with d
- Empirical error rate with the (false) intra-independent model better with d!



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Some alternatives for reducing variance

- Dimension reduction in non-canonical space (PCA-like typically)
- Dimension reduction in the canonical space (variable selection)
- Model parsimony in the initial HD space (constraints on model parameters)

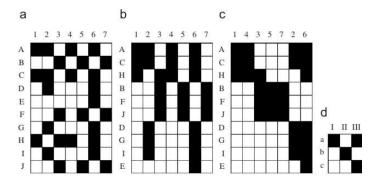
But which kind of parsimony?

- lacktriangle Remember that clustering is a way for dealing with large n
- Why not reusing this idea for large *d*?

Co-clustering

It performs parsimony of row clustering through variable clustering

From clustering to co-clustering



[Govaert, 2011]

Notations

- \mathbf{z}_i : the cluster of the row i
- **w**_j: the cluster of the column j
- **(** z_i , w_j): the block of the element x_{ij} (row i, column j)
- $\mathbf{z} = (\mathbf{z}_1, \dots, \mathbf{z}_n)$: partition of individuals in K custers of rows
- $\mathbf{w} = (\mathbf{w}_1, \dots, \mathbf{w}_d)$: partition of variables in L clusters of columns
- (z, w): bi-partition of the whole data set x
- lacksquare Both space partitions are respectively denoted by $\mathcal Z$ and $\mathcal W$

Restriction

All variables are of the same kind (research in progress for overcoming that...)

- Generalization of some existing non-probabilistic methods
- Extend the latent class principle of local (or conditional) independence
- Thus x_{ij} is assumed to be independent once z_i and w_j are fixed $(\alpha = (\alpha_{kl}))$:

$$p(\mathbf{x}|\mathbf{z},\mathbf{w};\alpha) = \prod_{i,j} p(x_{ij};\alpha_{z_iw_j})$$

- $m{\pi}=(\pi_k)$: vectors of proba. π_k that a row belongs to the kth row cluster
- $m{\rho}=(
 ho_k)$: vectors of proba. ho_k that a row belongs to the /th column cluster
- Independence between all z_i and w_j
- **E**xtension of the traditional mixture model-based clustering $(\alpha = (\alpha_{kl}))$:

$$p(\mathbf{x}; \boldsymbol{\theta}) = \sum_{(\mathbf{z}, \mathbf{w}) \in \mathcal{Z} \times \mathcal{W}} \prod_{i, i} \pi_{z_i} \rho_{w_j} p(x_{ij}; \boldsymbol{\alpha}_{z_i w_j})$$

Distribution for different kinds of data

[Govaert and Nadif, 2014] The pdf $p(\cdot; \alpha_{z_i w_j})$ depends on the kind of data x_{ij} :

- Binary data: $x_{ij} \in \{0,1\}$, $p(\cdot; \alpha_{kl}) = \mathcal{B}(\alpha_{kl})$
- Categorical data with *m* levels:

$$\mathbf{x}_{ij} = \{x_{ijh}\} \in \{0,1\}^m \text{ with } \sum_{h=1}^m x_{ijh} = 1 \text{ and } \mathbf{p}(\cdot; oldsymbol{lpha}_{kl}) = \mathcal{M}(oldsymbol{lpha}_{kl}) \text{ with } oldsymbol{lpha}_{kl} = \{lpha_{kjh}\}$$

- Count data: $x_i^j \in \mathbb{N}$, $p(\cdot; \alpha_{kl}) = \mathcal{P}(\mu_k \nu_l \gamma_{kl})^9$
- Continuous data: $x_i^j \in \mathbb{R}$, $p(\cdot; \alpha_{kl}) = \mathcal{N}(\mu_{kl}, \sigma_{kl}^2)$

⁹The Poisson parameter is here split into μ_k and ν_l the effects of the row k and the column l respectively and γ_{kl} the effect of the block kl. Unfortunately, this parameterization is not identifiable. It is therefore not possible to estimate simultaneously μ_k , ν_l and γ_{kl} without imposing further constraints. Constraints $\sum_k \pi_k \gamma_{kl} = \sum_l \rho_l \gamma_{kl} = 1 \text{ and } \sum_k \mu_k = 1, \sum_l \nu_l = 1 \text{ are a possibility.}$

Number of parameters

Extreme parsimony ability

Binary	$dim(oldsymbol{\pi}) + dim(oldsymbol{ ho}) + KL$
Categorical	$\dim(\pi) + \dim(\rho) + \mathit{KL}(m-1)$
Contingency	$\dim(\pi) + \dim(\rho) + KL$
Continuous	$\dim(\pi) + \dim(\rho) + 2KL$

Model

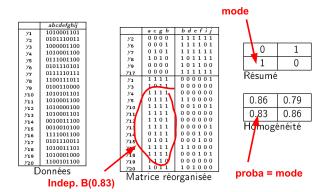
Very parsimonious so well suitable for the (ultra) HD setting

nb. param._{HD} = nb. param._{classic}
$$\times \frac{L}{d}$$

Other advantage: stay in the canonical space thus meaningful for the end-user

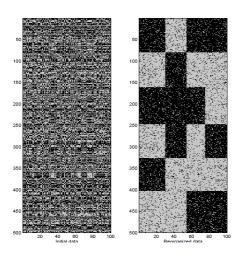
Binary illustration: easy interpretation

[Govaert, 2011]



Binary illustration: user-friendly visualization

[Govaert, 2011]



$$n = 500, d = 10, K = 6, L = 4$$

MLE estimation: log-likelihood(s)

- **Remember Lesson 3:** first estimate θ , then deduce estimate of (z, w)
- Observed log-likelihood: $\ell(\theta; \mathbf{x}) = \ln p(\mathbf{x}; \theta)$
- MLE:

$$\hat{oldsymbol{ heta}} = rg\max_{oldsymbol{ heta}} \ell(oldsymbol{ heta}; oldsymbol{x})$$

■ Complete log-likelihood:

$$\ell_c(\boldsymbol{\theta}; \mathbf{x}, \mathbf{z}, \mathbf{w}) = \ln p(\mathbf{x}, \mathbf{z}, \mathbf{w}; \boldsymbol{\theta})$$

$$= \sum_{i,k} z_{ik} \log \pi_k + \sum_{k,l} w_{jl} \log \rho_l + \sum_{i,j,k,l} z_{ik} w_{jl} \log p(x_i^j; \boldsymbol{\alpha}_{kl})$$

Be careful with asymptotics...

If $\ln(d)/n \to 0$, $\ln(n)/d \to 0$ when $n \to \infty$ and $d \to \infty$, then the MLE is consistent [Brault *et al.*, 2017]

MLE estimation: EM algorithm

■ E-step of EM (iteration q):

$$Q(\theta, \theta^{(q)}) = E[\ell_{c}(\theta; \mathbf{x}, \mathbf{z}, \mathbf{w}) | \mathbf{x}; \theta^{(q)}]$$

$$= \sum_{i,k} \underbrace{p(z_{i} = k | \mathbf{x}; \theta^{(q)})}_{t_{ik}^{(q)}} \ln \pi_{k} + \sum_{j,l} \underbrace{p(w_{i} = l | \mathbf{x}; \theta^{(q)})}_{s_{jl}^{(q)}} \ln \rho_{l}$$

$$+ \sum_{i,j,k,l} \underbrace{p(z_{i} = k, w_{j} = l | \mathbf{x}; \theta^{(q)})}_{e_{ijkl}^{(q)}} \ln p(x_{ij}; \alpha_{kl})$$

 \blacksquare M-step of EM (iteration q): classical. For instance, for the Bernoulli case, it gives

$$\pi_k^{(q+1)} = \frac{\sum_i t_{ik}^{(q)}}{n}, \quad \rho_l^{(q+1)} = \frac{\sum_j s_{jl}^{(q)}}{d}, \quad \alpha_{kl}^{(q+1)} = \frac{\sum_{i,j} e_{ijk}^{(q)} x_{ij}}{\sum_{i,j} e_{ikl}^{(q)}}$$

 $e_{iikl}^{(q)}$ is usually intractable. . .

- Consequence of dependency between \mathbf{x}_{ij} s (link between rows and columns)
- Involve K^nL^d calculus (number of possible blocks)
- Example: if n = d = 20 and K = L = 2 then 10^{12} blocks
- Example (cont'd): 33 years with a computer calculating 100,000 blocks/second

Alternatives to EM

■ Variational EM (numerical approx.): conditional independence assumption

$$p(z, w|x; \theta) \approx p(z|x; \theta)p(w|x; \theta)$$

- SEM-Gibbs (stochastic approx.): replace E-step by a S-step approx. by Gibbs
 - $z|x, w; \theta$ and $w|x, z; \theta$

MLE: variational EM (1/2)

- Use a general variational result from [Hathaway, 1985]
- Maximizing $\ell(\theta; \mathbf{x})$ on θ is equivalent to maximize $\tilde{\ell}_c(\theta; \mathbf{x}, \mathbf{e})$ on (θ, \mathbf{e})

$$\tilde{\ell}_c(\boldsymbol{\theta}; \boldsymbol{x}, \mathbf{e}) = \sum_{i,k} t_{ik} \ln \pi_k + \sum_{j,l} s_{jl} \ln \rho_l + \sum_{i,j,k,l} e_{ijkl} \ln \mathsf{p}(x_{ij}; \boldsymbol{\alpha}_{kl})$$

where
$$\mathbf{e}=(e_{ijkl})$$
, $e_{ijkl}\in\{0,1\}$, $\sum_{k,l}e_{ijkl}=1$, $t_{ik}=\sum_{j,l}e_{ijkl}$, $s_{jl}=\sum_{i,k}e_{ijkl}$

- Of course maximizing $\ell(\theta; \mathbf{x})$ or $\tilde{\ell}_c(\theta; \mathbf{x}, \mathbf{e})$ are both intractable
- Idea: restriction on **e** to obtain tractability $\mathbf{e}_{ijkl} = t_{ik} s_{jl}$
- New variables are thus now $\mathbf{t} = (t_{ik})$ and $\mathbf{s} = (s_{il})$
- As a consequence, it is a maximization of a lower bound of the max. likelihood

$$\max_{\boldsymbol{\theta}} \ell(\boldsymbol{\theta}; \boldsymbol{x}) \geq \max_{\boldsymbol{\theta}. \mathbf{t}. \mathbf{s}} \tilde{\ell}_c(\boldsymbol{\theta}; \boldsymbol{x}, \mathbf{e})$$

Approximated E-step

$$Q(\boldsymbol{\theta}, \boldsymbol{\theta}^{(q)}) \approx \sum_{i,k} t_{ik}^{(q)} \ln \pi_k + \sum_{j,l} s_{jl}^{(q)} \ln \rho_l + \sum_{i,j,k,l} t_{ik}^{(q)} s_{jl}^{(q)} \ln p(x_{ij}; \boldsymbol{\alpha}_{kl})$$

- We called it now VEM
- Also known as mean field approximation
- Consistency of the variational estimate [Brault et al., 2017]

MLE: local maxima

- More local maxima than in classical mixture models
- It is a consequence of many more latent variables (blocks)
- Thus: either many VEM runs, or use the SEM-Gibbs algorithm

MLE: SEM-Gibbs

■ We have already seen the SEM algorithm in Lesson 3 (thus we do not detail more)

- It limits dependency to starting point, so it limits local maxima
- The S-step: a draw $(\pmb{z}^{(q)}, \pmb{w}^{(q)}) \sim p(\pmb{z}, \pmb{w}|\pmb{x}; \pmb{\theta}^{(q)})$ instead an expectation
- But it is still intractable, thus use a Gibbs algorithm to approx. this draw

Approximated S-step

Two easy draws

$$\mathbf{z}^{(q)} \sim \mathsf{p}(\mathbf{z}|\mathbf{w}^{(q-1)},\mathbf{x};\mathbf{\theta}^{(q)})$$

and

$$\mathbf{w}^{(q)} \sim p(\mathbf{w}|\mathbf{z}^{(q)},\mathbf{x};\mathbf{\theta}^{(q)})$$

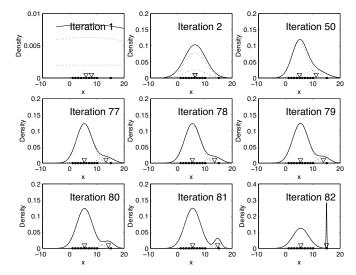
- Rigorously speaking, many draws within the S-step should be performed
- Indeed, Gibbs has to reach a stochastic convergence
- In practice it works well while saving computation time

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MLE: degeneracy

- More degenerate situations than in classical mixture models
- It is again a consequence of many more latent variables (blocks)
- The Bayesian regularization (instead MLE) can be an answer

Illustration of a degenerate situation



lacktriangle Everything passes by the posterior distribution of $oldsymbol{ heta}$

$$\mathsf{p}(heta|\mathbf{x}) \propto \underbrace{\mathsf{p}(\mathbf{x}| heta)}_{\mathsf{log-likelihood}} \underbrace{\mathsf{p}(heta)}_{\mathsf{prior}}$$

■ Then, take (for instance) the MAP as a θ estimate (use a VEM like algo...)

$$\hat{\boldsymbol{\theta}} = \arg\max_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\boldsymbol{x})$$

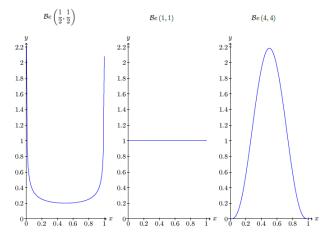
Bayesian estimation: limiting degeneracy

- Interest for avoiding degeneracy is the prior: it acts as a penalization term
- lacktriangle Typical choices are Dirichlet for π and ho (with independence between π , ho, lpha)

$$\mathsf{p}(heta) = \underbrace{\mathsf{p}(\pi)}_{D_K(a,\dots,a)} imes \underbrace{\mathsf{p}(
ho)}_{D_L(a,\dots,a)} imes \underbrace{\mathsf{p}(lpha)}_{\mathsf{model dependent}}$$

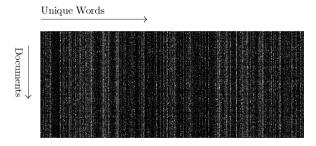
- The Dirichlet distribution is conjugate, thus easy calculus
- Control degeneracy frequency with the a value:
 - **a** = 1: uniform prior, so $\hat{\theta}$ is strictly the MLE (no regularisation)
 - a = 1/2: Jeffreys prior, classical (no informative prior) but may favor degeneracy
 - a = 4: a rule of thumb working well for limiting degeneracy frequency

Bayesian estimation: prior overview



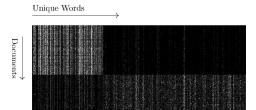
Document clustering (1/2)

- Mixture of 1033 medical summaries and 1398 aeronautics summaries
- Lines: 2431 documents
- Columns: present words (except stop), thus 9275 unique words
- Data matrix: cross counting document×words
- Poisson model



Document clustering (2/2)

$$p(\hat{\mathbf{z}} \neq \mathbf{z}) \leq 2n \exp\Big\{-\frac{1}{8}d\Big[\underbrace{\min_{k \neq k'}|\tau_k - \tau_{k'}|}_{\text{overlap}}\Big]\Big\} + K(1 - \min_k \pi_k)^n$$



Results with 2×2 blocs

	Medline	Cranfield
Medline	1033	0
Cranfield	0	1398

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Running BlockCluster

Configuration

If you change the configuration of your job and save it, it will start a new process with the updated parameters. This will erase previous results.

Parameters		
Title	Trial BlockCluster	
Data File	Blockcluster-Example.csv	
Data Type	Categorical ▼	θ
Rows Cluster Groups	1:5	•
Column Cluster Groups	1:5	θ
	Update	

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Running BlockCluster



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Running BlockCluster

Model	Criterion	Nb Clusters	Error	1
pik_rhal_multi	ICL (-45557:1)	[2,3]	No error	
plk_rhol_multi	ICL (-45563.3)	[3,3]	No error	
plk_rhol_multi	ICL (-45566.6)	[2,4]	No error	
plk_rhol_multi	ICL (-45573.9)	[4,3]	No error	
pik_rhol_multi	ICL (-45574.6)	[5,3]	No error	
pik_rhol_multi	ICL (-45577.7)	[3,4]	No error	
pik_rhol_multi	ICL (-45578.8)	[2,5]	No error	

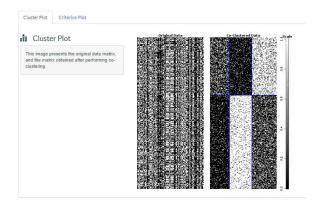
Cluster Plot

If Model Criterion

This chart represents the criterion value for the control of t

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Running BlockCluster



Outline

- 1 Introduction
- 2 Model-based clustering
- 3 Mixmod in MASSICCO
- 4 MixtComp in MASSICCC
- 5 BlockCluster in MASSICCC
- 6 Conclusion

- Use probabilistic modelling as a mathematical guideline
- Use the MASSICCC platform for user-friendly implementation

https://massiccc.lille.inria.fr/