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# Improving the inverse modeling of a trace isotope: how precisely can radium-228 fluxes toward the ocean and submarine groundwater discharge be estimated?

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Abstract. Radium-228 (228Ra), an almost conservative trace isotope in the ocean, supplied from the continental shelves and removed by a known radioactive decay ( $T_{1/2}$  = 5.75 years), can be used as a proxy to constrain shelf fluxes of other trace elements, such as nutrients, iron, or rare earth elements. In this study, we perform inverse modeling of a global <sup>228</sup>Ra dataset (including GEOSECS, TTO and GEO-TRACES programs, and, for the first time, data from the Arctic and around the Kerguelen Islands) to compute the total <sup>228</sup>Ra fluxes toward the ocean, using the ocean circulation obtained from the NEMO 3.6 model with a 2° resolution. We optimized the inverse calculation (source regions, cost function) and find a global estimate of the <sup>228</sup>Ra fluxes of  $8.01-8.49 \times 10^{23}$  atoms yr<sup>-1</sup>, more precise and around 20 % lower than previous estimates. The largest fluxes are in the western North Atlantic, the western Pacific and the Indian Ocean, with roughly two-thirds in the Indo-Pacific Basin. An estimate in the Arctic Ocean is provided for the first time  $(0.43-0.50\times10^{23} \text{ atoms yr}^{-1})$ . Local misfits between model and data in the Arctic, the Gulf Stream and the Kuroshio regions could result from flaws of the ocean circulation in these regions (resolution, atmospheric forcing). As radium is enriched in groundwater, a large part of the <sup>228</sup>Ra shelf sources comes from submarine groundwater discharge (SGD), a major but poorly known pathway for terrestrial mineral elements, including nutrients, to the ocean. In contrast to the <sup>228</sup>Ra budget, the global estimate of SGD is rather unconstrained, between 1.3 and  $14.7 \times 10^{13} \text{ m}^3 \text{ yr}^{-1}$ , due to high uncertainties on the other sources of <sup>228</sup>Ra, especially diffusion from continental shelf sediments. Better precision on SGD cannot be reached by inverse modeling until a proper way to separate the contributions of SGD and diffusive release from sediments at a global scale is found.

## 1 Introduction

Trace elements and isotopes (TEIs) are low-concentration components of the ocean, but they contain decisive information for our understanding of its dynamics. The international program GEOTRACES has been designed to improve our knowledge on the TEI concentrations and the oceanic processes controlling their distribution, by means of observations, modeling and laboratory experiments. Since 2006, GEOTRACES cruises have been mapping the global distribution of tens of these TEIs. Some of them are studied because they constitute micronutrients for living organisms, like iron (Fe), or are pollutants, like lead (Pb) and cadmium (Cd). Others are proxies of ocean dynamics or of biogeochemical processes. For instance, neodymium (Nd) is a proxy of the exchanges between seabed and seawater (Jeandel et al., 2007), and thorium 234 (<sup>234</sup>Th) is related to the biological carbon pump (Clegg and Whitfield, 1991; Buesseler et al., 1992; Henson et al., 2011). Radium (Ra) isotopes are of particular interest as they are proxies of all TEI fluxes from sediments and continents toward the ocean. More specifically, radium isotopes have been used to estimate a

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still poorly known pathway to the ocean: submarine ground-water discharge (SGD).

SGD is defined as the flux of water from coastal aquifers to the ocean, regardless of its composition and origin. Part of it is meteoric freshwater, but the largest part is infiltrated seawater (Burnett et al., 2006; Moore, 2010b). In these aquifers, water gets enriched in nutrients and trace elements, released from soils and rocks or coming from land pollution, before flowing back to the ocean. It has traditionally been considered that the main coastal sources of terrestrial mineral elements to the ocean are rivers, but there is growing evidence that SGD is in fact a source of nutrients of the same order of magnitude, affecting the biogeochemistry at all scales, from coastal regions (Hwang et al., 2005; Kim et al., 2005) to ocean basins (Moore et al., 2008; Rodellas et al., 2015). SGD is a suspected cause of algal blooms, including harmful algal blooms (LaRoche et al., 1997), and a pathway for contamination. In spite of this, its contribution is still much less precisely known than the river inputs. Because of the strong heterogeneity in its distribution and intensity, properly estimating SGD by direct methods requires an intense sampling, which is far from being fulfilled (Burnett et al., 2006; Moore, 2010a). Indirect methods should then be used. Radium isotopes offer a great potential due to their relation to SGD and their simple chemistry.

All the four natural radium isotopes,  $^{223}$ Ra ( $T_{1/2} = 11.4 \,\text{day}$ ),  $^{224}$ Ra ( $T_{1/2} = 3.6 \,\text{d}$ ),  $^{226}$ Ra ( $T_{1/2} = 1602 \,\text{years}$ ), and  $^{228}$ Ra ( $T_{1/2} = 5.75$  years), are produced within the rocks by the radioactive decay of thorium. Since radium is far more soluble in water than thorium, its main source is not the decay of dissolved thorium in the ocean, but dissolution from lithogenic material. Therefore it is used as a tracer of boundary fluxes. This element is released into the ocean by three main sources located on the continental shelf: desorption from riverine particles, diffusion from seabed sediments, and SGD, which are highly enriched. Dust inputs and dissolved riverine <sup>228</sup>Ra each account for less than 1 % of all inputs (Moore et al., 2008). In the ocean, Ra is almost conservative. It is removed by radioactive decay and scavenging. Scavenging is associated with a residence time of approximately 500 years (Moore and Dymond, 1991), making it negligible for all Ra isotopes except <sup>226</sup>Ra. Then the distributions of the other three isotopes, whose radioactive sinks are known, depend only on the source distribution and transport by the ocean circulation. Since their half-life timescales are small relative to the timescale of basin-wide horizontal mixing, typically a few years to a few decades, <sup>223</sup>Ra and <sup>224</sup>Ra are unsuitable for large-scale studies. Radium-228, whose half-life of 5.75 years is short enough to neglect scavenging but long enough to consider the global ocean, is suitable for global-scale analyses, and is thus considered here.

A simple way to use the information provided by this isotope is to make an observation-based inventory of the ocean <sup>228</sup>Ra. At steady state, the supply of <sup>228</sup>Ra must balance the loss from disintegration – i.e., 12 % every year. Accord-

ing to Charette et al. (2016), these total <sup>228</sup>Ra fluxes can be used to estimate fluxes of nutrients, iron, and rare earth elements. Radium-228 fluxes are also a way to estimate the SGD by subtracting the contribution from rivers and the release from sediments by diffusion and bioturbation, and dividing the remaining flux by the mean <sup>228</sup>Ra concentration in groundwater. By this method SGD has been estimated at  $0.03-0.48 \times 10^{13} \,\mathrm{m}^3 \,\mathrm{yr}^{-1}$  in the Mediterranean Sea (Rodellas et al., 2015) and  $2-4 \times 10^{13}$  m<sup>3</sup> yr<sup>-1</sup> in the Atlantic Ocean (Moore et al., 2008). These direct approaches suffer from strong potential biases when the data are sparse, because the total amount of <sup>228</sup>Ra in the ocean is then roughly estimated, using regional averages (Rodellas et al., 2015) of observations or linear interpolations (Moore et al., 2008). Therefore, they are suitable only in regions with dense sampling, such as the Atlantic Basin.

Inverse modeling techniques represent an alternative and powerful approach to estimate the fluxes, providing that the ocean circulation is known with sufficient accuracy. Inverse modeling is based on three elements: observations of the physical quantity of interest (here oceanic <sup>228</sup>Ra concentration), a forward prognostic model simulating the same quantity as a function of the unknown variables to be assessed (here the shelf fluxes), and an algorithm finding the values of the unknowns (fluxes) minimizing the misfit between model and observations. The advantage of this method is that no arbitrary averaging or interpolation is required: observed concentrations are simply compared to model concentrations at the same points. It is expected to be robust and consistent, since the model is based on physical considerations. Kwon et al. (2014) used such an inverse modeling approach to produce a global estimate of <sup>228</sup>Ra fluxes of 9.1–  $10.1 \times 10^{23}$  atoms yr<sup>-1</sup> between 60° S and 70° N, corresponding to  $9-15 \times 10^{13}$  m<sup>3</sup> yr<sup>-1</sup> of SGD. Their study is based on a data-constrained global ocean circulation model (DeVries and Primeau, 2011), considering 50 source regions on the continental shelf, and minimizing an ordinary least-squares cost function.

In this study, we estimate the radium fluxes from all continental shelves around the world and localize the most intense sources, using an inverse modeling technique with more data than previous studies (Kwon et al., 2014). The dataset has been augmented with data from two recent GEOTRACES cruises and from the Southern Ocean, the North Pacific, the Mediterranean Sea, and the Indonesian seas. It also contains data from the Arctic, a basin absent from Kwon's study. The forward model is built on the ocean general circulation model (OGCM) NEMO (Nucleus for European Modeling of the Ocean). Our main improvement is a careful analysis of sensitivity and errors, which reveals that the result depends on the model mathematical parameters, such as the cost function and the number of regional sources that are considered. We have performed several inversions and analyzed the residuals and uncertainties, to determine the most appropriate mathematical parameters and evaluate the precision of the flux estimates. This paper is organized as follows. In Sect. 2, we describe the different aspects of the inversion technique – e.g., the global dataset, the forward model based on the NEMO OGCM, the different cost functions, the choice of the source regions related to SGDs, and the inverse method. Section 3 presents the main results – e.g., the global and regional estimates of <sup>228</sup>Ra supply, and the sensitivity of these estimates to several parameters of our approach, such as the cost function or the number of coastal <sup>228</sup>Ra sources. Section 4 compares our results with results obtained in previous studies, and discusses issues associated with such an approach, with an emphasis on SGD.

#### 2 Methods

# 2.1 <sup>228</sup>Ra dataset

Since the late 1960s (Moore, 1969; Kaufman et al., 1973) tens of oceanographic cruises have carried out measurements of <sup>228</sup>Ra (e.g., articles listed in Table S1 in the Supplement). The dataset used in this study includes, among others, observations from three international programs sampling trace elements. Data from the Indian Ocean cruise of GEOSECS (GEochemical Ocean SECtions Study), in 1977-1978, are included. From 1981 to 1989, the TTO (Transient Tracers in the Oceans) program produced a considerable number of <sup>228</sup>Ra measurements in the Atlantic, from 80° N to 60° S, at all depths, making the Atlantic Ocean the best sampled ocean by far. Currently, new data from all oceans are being produced by GEOTRACES. In total, 6059 data from all basins are used, of which 2789 are shallower than 10 m, 1107 are located between 10 and 200 m deep, 606 between 200 and 600 m deep, and 1557 deeper than 600 m. Our dataset comprises 1359 more measurements than in Kwon et al. (2014): two GEOTRACES sections, GA03 (United States-Cape Verde-Portugal) and GP16 (Ecuador-French Polynesia), and data from the Arctic, the Southern Ocean, the North Pacific, the Mediterranean Sea, and the Indonesian seas have been added in the present study (see Table S1). These supplementary observations extend the data coverage to regions north of 70° N and around the Kerguelen Islands. In a near future, other sections from the GEO-TRACES program will complete the global covering and may help in studying deeper sources.

For the purpose of our study, data have been averaged in each model grid cell (see Sect. 2.2 for more details on the model configuration), leading to the 3076 cell averages shown in Fig. 1. The density of the measurement is noticeably uneven. The Atlantic Ocean, north of 20° S, and the Arctic, are the most densely covered basins and the only regions with a significant number of data at depth deeper than 10 m. Other regions are sparsely sampled, leaving wide areas with no or very few measurements, like the western Indian Ocean,

the equatorial Pacific, or the Pacific sector of the Southern Ocean

Data are expressed in concentration or activity units, with the following conversion factor:  $1 \, \text{dpm} \, \text{m}^{-3} = 4.36 \times 10^6 \, \text{atoms} \, \text{m}^{-3}$ . They range from 0.04 to 724.5 dpm m<sup>-3</sup>. The highest concentrations are found in the Bay of Bengal and the coastal seas of eastern Asia, the lowest values are located in the Southern Ocean. Concentrations are generally higher than  $10 \, \text{dpm} \, \text{m}^{-3}$  in the Indian Ocean, the Atlantic, and the Pacific north of  $30^\circ$  N, lower in the rest of the Pacific. In the Atlantic, west of a line running from the Amazon Delta to Newfoundland, most concentrations are higher than  $30 \, \text{dpm} \, \text{m}^{-3}$ .

#### 2.2 Forward tracer model

The second requirement of the inversion technique is a  $^{228}$ Ra transport model, allowing to link in situ observations to the boundary conditions or shelf sources of  $^{228}$ Ra. The transport equation of tracer  $A_i$  (originating from the source region i) is given by

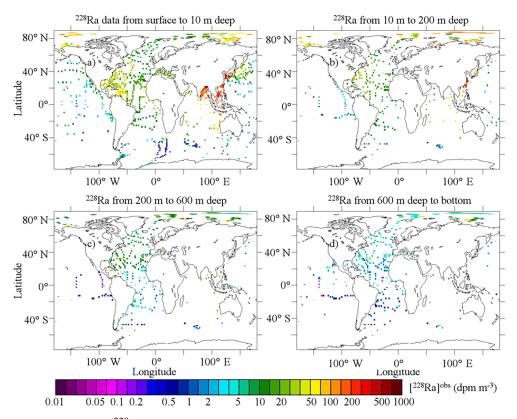
$$\frac{\partial A_i}{\partial t} = -\boldsymbol{U} \cdot \nabla A_i + \nabla \cdot (\boldsymbol{K} \nabla A_i) - \lambda A_i + S_i. \tag{1}$$

U and K are the velocity field and the eddy diffusivity coefficient respectively. The two first terms on the right together constitute transport. They are derived from NEMO 3.6 model using OPA (Madec, 2015) as a general circulation component, coupled with the sea-ice model LIM3 (Vancoppenolle et al., 2009), with an ORCA2 global configuration. The model has a horizontal resolution of  $2^{\circ} \times 2^{\circ} \cos \phi$  (where  $\phi$  is the latitude) enhanced to  $0.5^{\circ}$  near the equator. The mesh is tripolar in order to overcome singularities, the North Pole being replaced by two inland poles in the Northern Hemisphere. It has 31 vertical levels, ranging from the surface to 6000 m deep, the upper layer covering the first 10 m. The simulation is forced by a seasonal climatological dataset, based on NCEP/NCAR reanalysis and satellite data.

 $\lambda$  is the radioactive decay constant,  $0.12\,\mathrm{yr^{-1}}$ , given by the half-life of  $^{228}$ Ra which is 5.75 years. Decay is the sole sink. It is known and does not depend on environmental parameters, leaving the source term as the only unknown to be determined by the inversion technique.

 $S_i$  is the source term specific to the ith region, representing riverine inputs, sedimentary diffusive fluxes and groundwater discharge fluxes of  $^{228}$ Ra from region i. In this study, sources are assumed to be only on the continental shelf, defined as the seabed shallower than 200 m. This depth range, spanning 16 model levels, is chosen because it is where most groundwater discharge outflows. However, diffusion from sediments also occurs at larger depths (Hammond et al., 1990).

Equation (1) shows that the  $A_i$  fields depend linearly on  $S_i$ . That means that any <sup>228</sup>Ra distribution can be written as a linear combination of the  $A_i$  fields. It is important to emphasize that the circulation is supposed to be "perfect" –



**Figure 1.** Observed concentrations of <sup>228</sup>Ra in the global ocean, averaged when available on the ORCA2 cells used for inversion. Data plotted on subfigures (**b**, **c**, **d**) are also vertically averaged in order to take all layers into account.

e.g., no correction of the U field is looked for. Nevertheless, the simulated circulation obviously suffers from deficiencies, and that point has to be kept in mind when interpreting the results. From now, what we refer to as "model concentration", [ $^{228}$ Ra] $^{mod}$ , is a linear combination of the tracer final concentrations  $A_i$ , the coefficients being the source intensities  $x_i$ :

$$[^{228}\text{Ra}]^{\text{mod}} = \sum_{i=1}^{n} A_i x_i.$$
 (2)

A non-optimal model concentration was computed by assuming a uniform constant flux per unit of surface everywhere, with a global fit using the average concentration estimated from the observations: this defines the first guess estimate before the inversion. The inversion undertaken in this study aims at optimizing the parameters  $x_i$  in order to minimize the total difference between the observations and the model  $^{228}$ Ra distribution.

As a consequence of its coarse horizontal resolution, continental shelves are only poorly resolved by the ORCA2 grid. The emitting surface is underestimated and some regions with narrow continental shelves would be completely omitted. To overcome that deficiency, sub-model grid-scale bathymetric variations are accounted for by comparing the model grid to a global 2' resolution bathymetry ETOPO2 of the National Geophysical Data Center (NGDC). The al-

gorithm is detailed in Aumont and Bopp (2006). According to this method, the total surface of continental shelf is  $2.73 \times 10^{13}$  m<sup>2</sup>, 73 % higher than the  $1.58 \times 10^{13}$  m<sup>2</sup> obtained with the coarser bathymetry.

The ocean-continent interface, including the Arctic and the Antarctic, is divided into 38 regions (Fig. 2). This first guess takes into account the sampling coverage (very low in the Antarctic for instance, and higher in the North Atlantic Basin or Bay of Bengal) and differences in the tracer distributions  $A_i$ , which should be large enough to give independent information. Delimitation is done by trial and error, using the posterior covariance matrix of the inversion (see Sect. 2.3): the number of sources is minimized by merging regions associated with negligible fluxes with close highly correlated regions. Most islands are ignored, because the areas of their continental shelf and thus their expected contributions to the <sup>228</sup>Ra balance are small. In the inversion process, islands can give rise to spurious fluxes to accommodate for other types of errors. The only islands considered in this study are the Kerguelen and Crozet islands, in the Southern Ocean, because many samples have been taken in their surroundings which make it possible to constrain their contributions. Because of the lack of measurements and the coarse model resolution, the Persian Gulf, the Red Sea, the Baltic Sea, the North Sea, and the Hudson Bay are not taken into account. The flux per

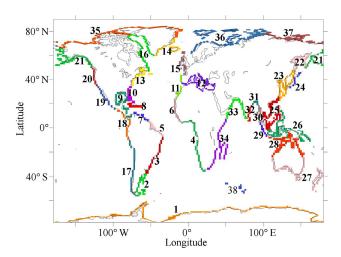


Figure 2. Model <sup>228</sup>Ra source regions.

unit of surface is assumed to be constant on each of the 38 emitting regions. Model simulations last for the equivalent of 100 years in order to reach a quasi-steady state. It is more than 17 times larger than the half-life of <sup>228</sup>Ra, so that the total amount of <sup>228</sup>Ra varies by less than 0.001%. As there are not enough data to study global interannual or seasonal variations, we do not take seasonal variations of <sup>228</sup>Ra concentrations into account. We implicitly assume that radium concentration is constant over time, and work with average concentrations over the 100th year of simulation.

#### 2.3 Inverse method

The last requirement of the inversion technique is to define a cost function measuring the misfit between the data and the model. This cost is then minimized by a method already used to assess air–sea gas fluxes (Gloor et al., 2001; Mikaloff Fletcher et al., 2006; Jacobson et al., 2007) and oceanic heat fluxes (Resplandy et al., 2016). If  $\eta$  represents the residuals, namely the discrepancy between model results and observations:

$$\mathbf{A}\mathbf{x} = [^{228}\text{Ra}]^{\text{obs}} + \boldsymbol{\eta}. \tag{3}$$

Both sides are vectors, representing radium concentrations in each model cell containing data. **A** is the matrix of footprints, representing the circulation model and composed of the  $A_i$  at each data point. x is the vector of unknowns, the flux of  $^{228}$ Ra per unit of surface of each region. As only shelf sources are modeled and as data coverage below 200 m is sparse, only data shallower than 200 m are considered.

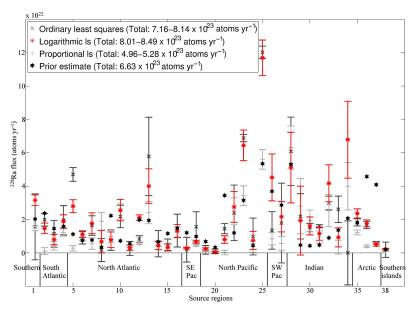
The distance between model concentrations and data is summed up in a scalar, the cost function C(x). We look for the optimal flux vector  $x_{\text{opt}}$  minimizing the cost function. Different choices for C are possible, depending on the assumed probability distribution for the prior error on data and model footprints A. Errors due to biased sampling are not

considered here. All errors are supposed to be uncorrelated. In this study, three different cost functions have been tested and minimized. They all correspond to the sum of squares of a specific type of residuals. Their respective equations are listed in Table 1. The first one is an ordinary least-squares cost function, Cols. According to the Gauss-Markov theorem, its minimization produces the best linear unbiased estimator when prior errors have no correlation, zero expectation, and the same variance. It is the simplest least-squares method, chosen in the study by Kwon et al. (2014). This function gives the same weight to all observations. However, the hypothesis of the homogeneity of the variance is questionable: far offshore, where concentrations are lower and less sensitive to small changes in coastal sources, observations and model errors can be expected to be lower. Neglecting this fact means these data are not fully exploited, as their contribution to the cost function is relatively small. Two other cost functions with a higher weight for smaller values are then considered for comparison. They are assuming heteroscedastic data, with higher variances for higher concentrations. In the proportional least-squares cost function  $C_{prp}$ , the error standard deviation is supposed to be proportional to the observed concentrations. The logarithmic least-squares cost function  $C_{log}$  works differently and assumes the logarithms of concentrations have the same error variance. It is less sensitive than  $C_{prp}$  to model overestimations and more to underestimations. It is the only cost function which is not a quadratic function of the sources.

The residuals after inversion indicate what the inverse model cannot fit. In a "perfect" inversion, these residuals should be assimilated to noise – e.g., small and without structure, due for instance to coarse resolution. In most inversions, that is not the case, and the distribution of residuals emphasizes biases or errors either in the chosen hypotheses (such as a perfect circulation) or in the setting of the inversion (number and choice of the regions). The posterior uncertainties on radium fluxes and correlations between regions are computed following the method described in Appendix A. A regional flux has a large uncertainty when it is constrained by few data or correlated to other regions (Gloor et al., 2001), and two regions are strongly correlated when the <sup>228</sup>Ra emitted by each is transported to the same places and are then harder to differentiate. The computed posterior uncertainties are precise only if all the preceding assumptions on prior errors are correct. The coherence of error assumptions with the results has to be checked (see Sect. 3.2). Along with the main inversion considering 38 regions, we performed four other inversions with a higher or lower number of regions in order to estimate the sensitivity to this parameter (see Sect. 3.3).

Table 1. Cost functions.

Cost function	Formula	Residuals
Ordinary least squares (ols)	$C_{\text{ols}}(x) = \sum_{j=1}^{p} ([^{228}\text{Ra}]_i^{\text{mod}} - [^{228}\text{Ra}]_i^{\text{obs}})^2$	$\eta_{\text{ols}}(x) = [^{228}\text{Ra}]^{\text{mod}} - [^{228}\text{Ra}]^{\text{obs}}$
Logarithmic least squares (log)	$C_{\log}(x) = \sum_{j=1}^{p} (\log[^{228}\text{Ra}]_i^{\text{mod}} - \log[^{228}\text{Ra}]_i^{\text{obs}})^2$	$ \eta_{\log}(x) = \log[^{228}\text{Ra}]^{\text{mod}} - \log[^{228}\text{Ra}]^{\text{obs}} $
Proportional least squares (prp)	$C_{\text{prp}}(x) = \sum_{j=1}^{p} \left( \frac{[^{228}\text{Ral}_{i}^{\text{mod}} - [^{228}\text{Ral}_{i}^{\text{obs}}]}{[^{228}\text{Ral}_{i}^{\text{obs}}]} \right)^{2}$	$ \eta_{\text{prp}}(x) = \frac{[^{228}\text{Ra}]^{\text{mod}} - [^{228}\text{Ra}]^{\text{obs}}}{[^{228}\text{Ra}]^{\text{obs}}} $



**Figure 3.** <sup>228</sup>Ra annual flux from each source within one standard deviation, after minimization of three cost functions. Prior estimates, proportional to shelf surfaces, are shown for comparison.

#### 3 Results

### 3.1 <sup>228</sup>Ra fluxes

The <sup>228</sup>Ra fluxes from each of the 38 regions, deduced by minimizing each of the three cost functions, are shown in Fig. 3 with their confidence intervals, and compared with the prior estimates. The global fluxes for each method are also shown. As they are sums of local fluxes, their standard deviations are proportionally lower.

The global  $^{228}$ Ra flux within 1 standard deviation is 8.01–  $8.49 \times 10^{23}$  atoms yr $^{-1}$  according to the  $C_{log}$  inversion. As we will explain in Sect. 3.2, this estimate is the most accurate of the three. Fluxes are found to be comparatively high in the North Atlantic (regions 5 to 16), in the western Pacific (22 to 27), and in the Indian Ocean (28 to 34), together accounting for 62.6% of the continental shelf and 82.8% of the global flux of  $^{228}$ Ra. The highest fluxes are located in the China seas (23 and 25) and in the eastern Indian Ocean (29 to 32), where the inversion process produces the largest in-

crease compared to prior estimates, and to a lesser extent on the east coast of North America (10 and 13). Conversely, inversion significantly reduces the prior estimates in the Arctic Ocean (35 to 37), in the Bering Sea (21), and in the eastern Pacific Ocean (17 to 20). Fluxes are also quite low in the Southern Ocean (1 and 38) and in the South Atlantic (2 to 4). The newly estimated Arctic and Antarctic sources are in the range of 0.43 to 0.50 and 0.31 to 0.37  $\times$  10<sup>23</sup> atoms yr $^{-1}$  respectively, accounting for 5 to 6.2 % and 3.6 to 4.6 % of the total sources.

Although having the same order of magnitude, uncertainties are generally lower than fluxes. They are highest in the western Pacific and Indian oceans (regions 22 to 34), because of data sparsity. It is lower in better sampled oceans: the Arctic (35 to 37) and the Atlantic (2 to 16) oceans, except for region 13 (Cape Hatteras to Newfoundland). The eastern Pacific (17 to 21) also has low uncertainties in absolute values, probably due to the low concentrations and prior errors there.

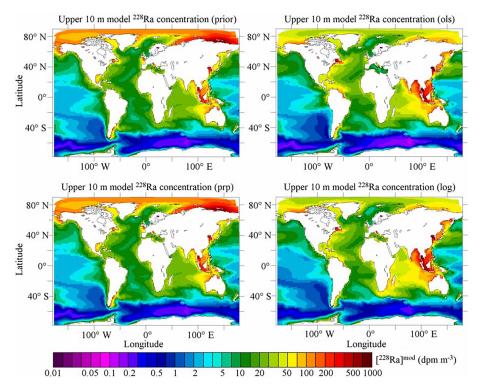


Figure 4. Model surface <sup>228</sup>Ra concentration after minimization of three cost functions. Prior estimate is shown for comparison.

The two other inversions produce roughly similar results, although fluxes are generally lower when derived from  $C_{\text{prp}}$ and generally have higher uncertainties when derived from  $C_{\rm ols}$ . The global <sup>228</sup>Ra flux is estimated to 7.16–8.14 ×  $10^{23} \text{ atoms yr}^{-1} \text{ with } C_{\text{ols}} \text{ and } 4.96-5.28 \times 10^{23} \text{ atoms yr}^{-1}$ with  $C_{\text{prp}}$ . The three inversions agree on which basins and continents have the largest and smallest sources. Yet, local disagreements occur. Regions 5 (Amazon delta), 21 (Alaska and Bering Sea) and 33 (Arabian Sea) have higher fluxes with  $C_{\rm ols}$  than with  $C_{\rm log}$ , with non-overlapping confidence intervals, whereas the contrary is true for regions 12 (Mediterranean), 26 (Indonesian seas) and 34 (East Africa).  $C_{prp}$ fluxes are generally lower, and in eight regions their confidence intervals overlap with none of the other inversions. Discrepancies happen because the fitting of the model to each observation implies different and possibly opposite effects on the source intensities. Each inversion uses different weights, which translates into different flux corrections. Fluxes from each of the three inversions are most dissimilar for regions where observations impose most dissimilar constraints, where the model fails to reconstruct the pattern of the data and has to choose between fitting some data or others in priority. In such cases, all the results should be considered carefully. When the confidence intervals between the different inversion techniques fail to overlap, it is likely that one or several estimates are incorrect. As algorithms are built by assuming a prior error statistics, it is likely that some rely on wrong assumptions.

**Table 2.** Correlation between data and model concentration fields.

Model	Linear correlation	Logarithmic correlation
Prior	0.383	0.809
ols	0.813	0.883
log	0.797	0.902
prp	0.754	0.877

#### 3.2 Model concentrations and residuals

The residuals – i.e., the differences between model concentrations and observations – determine how well the model reproduces the observations and quantify the improvements in the tracer distribution provided by the inversion. They are also a basic tool to identify biases in the model and to assess the quality of the assumptions.

Radium fluxes obtained by the inverse method largely improve the model match to observations compared to the prior radium flux (Figs. 4 and 5). The improvement is quantified by the increase in the model—data correlations (Table 2) and the decrease in the root mean square of the residuals (Table 3), a proxy of the cost function. The correlation coefficient is increased from 0.383 to 0.813 on a linear scale and from 0.809 to 0.902 on a logarithmic scale. The correlation is higher on a logarithmic scale because it is less sensitive to

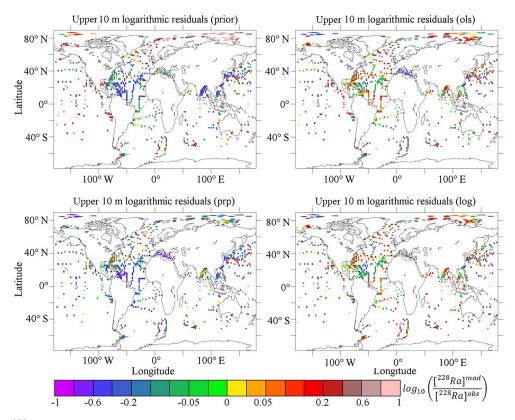


Figure 5. Surface <sup>228</sup>Ra logarithmic residuals after minimization of three cost functions. Prior estimate is shown for comparison.

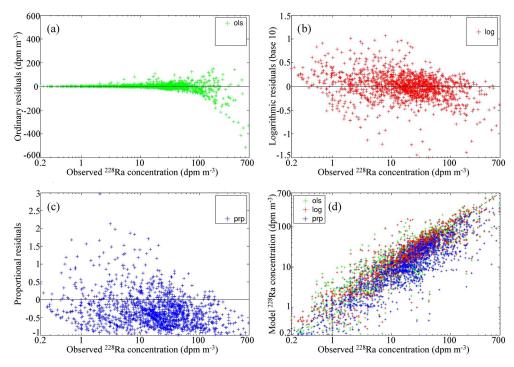
Table 3. Root mean square of residuals before and after inversion

Model	Ordinary residuals rms $(dpm m^{-3})$	Logarithmic residuals rms (no unit)	Proportional residuals rms (no unit)
Prior	70.7	0.892	1.91
ols	36.6	0.703	1.10
log	37.9	0.636	1.02
prp	47.4	0.940	0.558

the few very high residuals associated with the highest concentrations (see Fig. 6). On average, the inversion is able to reduce the ordinary residuals ( $C_{\rm ols}$ ) by a factor 2, logarithmic residuals ( $C_{\rm log}$ ) by 1.4, and proportional residuals ( $C_{\rm prp}$ ) by 3.5.

In spite of being smaller, the order of magnitude of the residuals remains comparable to the data (Fig. 5). On the one hand, in all oceans, positive and negative residuals are observed with no clear patterns at the scale of a few grid cells. Because of the rather low model resolution (2°, which is not sufficient to reproduce medium- and small-scale processes) and of issues associated with temporal (seasonal or higher frequency) variability of the data, this kind of "noise" is ex-

pected. It is consistent with the assumption of independent errors used when computing the error variances on fluxes. On the other hand, in several regions, residuals can display coherent large-scale patterns, which cannot be attributed to noise. These areas may suffer from systematic overestimations, like in the Gulf Stream region, the western Pacific between 20 and 40° N, and off eastern Siberia, or underestimations, such as in the center of the North Atlantic Gyre. These residuals point out to possible flaws in the model circulation. For instance, <sup>228</sup>Ra is quite homogeneously distributed in the western North Atlantic according to data, but in the model, the gradients are stronger, and no combination of sources manages to reduce these gradients. This is probably caused by a bias in the Atlantic circulation, with a too low exchange rate between inshore and offshore waters. In a 2° resolution model, mesoscale eddies are not represented and cannot transport <sup>228</sup>Ra south and east of the Gulf Stream or north of the North Brazil Current. Inversions minimize the misfit by increasing the fluxes from regions 5 (northern Brazil), 8 (Caribbean), and 10 (southern east coast of the US), making the model concentration too high close to the coast while still too low in the gyre. Such large-scale biases are not consistent with the assumption of no prior error correlation, which may lead to underestimation of flux uncertainties around these basins.



**Figure 6.** (a) Residuals, (b) logarithmic residuals, (c) proportional residuals, and (d) model <sup>228</sup>Ra concentration as a function of observed <sup>228</sup>Ra concentration. Lines of zero residuals are drawn in black.

Having assumed specific prior error statistics when choosing the cost functions, we need to check that there is no a posteriori contradiction. Figure 6 displays the residuals and model concentrations as a function of the observations. If the residuals depend on the observed concentrations, it means some observations are more precise than others, contain more information, and should be given a higher weight in order to obtain the best linear unbiased estimate. Figure 7 shows the probability density functions (PDFs) of residuals after all three inversions and compares them with a Gaussian curve representing the expected distribution given the root mean square of residuals. Each PDF should look like a Gaussian curve for the computed posterior uncertainties to be relevant descriptors of errors.

Figure 7a emphasizes that the ordinary residuals do not follow a Gaussian distribution. On the contrary, most residuals are very close to zero, while a small number of them are much higher than the standard deviation. Figure 6a shows that these high residuals occur at high concentrations only, and that error variance is not homogeneously distributed. Then high and low concentrations should not be given the same weights, as in  $C_{\rm ols}$ , but the highest concentrations should be given the lowest weights, as in  $C_{\rm log}$  and  $C_{\rm prp}$ . Flux estimates based on  $C_{\rm ols}$  are biased because the cost function puts more emphasis on high concentrations, and this method then tries to fit more specifically the misfits at high concentrations.  $C_{\rm ols}$  also produces very large error bars because the error variance is assumed to be constant and its computed

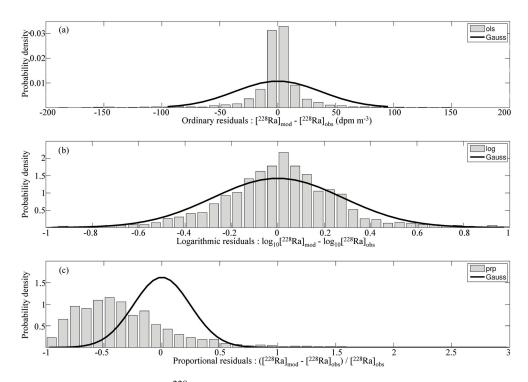
value, influenced by a few very large residuals, is larger than the actual error variances of the vast majority of data. Figures 6c and 7c show that the proportional residuals are not normally distributed either. They are not even symmetrical, as they cannot be smaller than -1 but they do not have an upper limit. This asymmetry produces a bias in the flux estimate. The algorithm based on  $C_{prp}$  is more sensitive to positive residuals because underestimations are never associated to proportional residuals lower than -1 whereas overestimations can produce residuals higher than 1. As a consequence, this method tends to reduce the fluxes. The hypothesis of constant variance is more realistic although the highest residuals occur at low concentrations. Finally, the distribution of the logarithmic residuals displayed in Figs. 6b and 7b is much closer to a Gaussian curve and much less dependent on concentrations, which makes the logarithmic cost function more relevant for this study.

## 3.3 Sensitivity to the number of regions

The choice of the regions (and their number) has been made rather subjectively, although several criteria have been used (spatial distribution of the observations, independence of the  $A_i$  fields). The global <sup>228</sup>Ra should ideally not depend upon the number of regions. Therefore, alternative region geometries have been tested for comparison. The <sup>228</sup>Ra fluxes are shown on Table 4 and the root mean square of their residuals is presented on Table 5. Case 1 inversion uses 52 regions: it was the original distribution of regions before some of them

Cost function	Case 1: 52 regions	Case 2: 38 regions	Case 3: 19 regions	Case 4: 12 regions	Case 5: 7 regions
ols	7.81–9.93	7.16-8.14	7.51–8.49	7.36–8.16	8.98–9.78
log	8.09–8.61	8.01-8.49	7.90–8.34	7.85–8.27	7.93-8.37
DID	5.15-5.47	4.96-5.28	4.70-4.98	4.25-4.49	3.75-4.01

**Table 4.** Model global  $^{228}$ Ra fluxes ( $10^{23}$  atoms yr $^{-1}$ ) with different numbers of source regions. The standard case with 38 regions is in bold.



**Figure 7.** Probability density functions (PDFs) of <sup>228</sup>Ra (a) ordinary residuals, (b) logarithmic residuals and (c) proportional residuals after inversion, compared with a Gaussian PDF (black full line) based on their standard deviations.

were merged to define the 38 standard regions of this study. It includes more regions in undersampled areas such as the Arctic, the South Atlantic, the western Indian, or the equatorial Pacific oceans. Case 5 has just one emitting region for each of the following ocean basins: Southern, South Atlantic, North Atlantic, South Pacific, North Pacific, Indian and Arctic. Case 3 and Case 4 are intermediate cases with source regions built by merging regions from Case 2.

The root mean square of residuals is a proxy of the cost function. On the one hand, this parameter should be as low as possible. Increasing the number of regions always decreases it because the number of degrees of freedom increases, which tends to improve the fit to the observations. In this inversion, the largest decrease is found between 7 and 12 source regions. Further increases in the number of regions have smaller impacts. On the other hand, too many source regions may produce spurious results. Some regional fluxes, with too few observations nearby to constrain them, would be computed using observations farther away, already used

by other fluxes. Because of the lower sensitivity of the concentrations at these farther-off locations, this process can create extreme fluxes, positive or even negative. The presence of physically impossible negative values, set to zero by the constraint of positivity, necessarily means such poor constraints exist. When 52 fluxes are computed, 5 to 7 of them, according to the cost function, are so poorly constrained that their fluxes have been set to zero to prevent them from being negative. This number is reduced to 1 with  $C_{ols}$  and zero with  $C_{\text{log}}$  and  $C_{\text{prp}}$  when there are 38 regions, and completely disappears with 19 or fewer regions. Regions with fluxes within the error estimate are also very poorly constrained by the observations and the circulation model: their number is also reduced from 26 out of 52 to 11 out of 38 with  $C_{\rm ols}$ , from 13 to 3 with  $C_{prp}$  and from 9 to 3 with  $C_{log}$ . All these fluxes have a low impact on the cost function, but make the global <sup>228</sup>Ra flux less precise. This analysis shows that Case 1 (with 52 regions) is not constrained enough and that Cases 3 to 5

Cost function Case 1: 52 regions Case 2: 38 regions Case 3: 19 regions Case 4: 12 regions Case 5: 7 regions ols 35.8 36.6 37.6 38.2 42.9 0.623 0.636 0.656 0.671 0.781 log 0.545 0.558 0.581 0.597 0.665 prp

**Table 5.** Root mean square of residuals after inversions with different numbers of source regions. The standard case with 38 regions is in bold.

display too large residuals. Therefore, Case 2 (38 regions) is considered to be the best choice.

The  $C_{\rm ols}$ -based global <sup>228</sup>Ra flux is varying in a nonmonotonic way, with a difference of 23 % between the highest and the lowest. The flux with 38 regions is lower than the fluxes computed with either more or fewer regions. This high sensitivity may be related to the high uncertainty associated with this inversion, certainly linked to the relatively poor data coverage. All the confidence intervals within one standard deviation from Case 1 to Case 4 overlap. The  $C_{prp}$ based global fluxes are always lower than the other fluxes, and they decrease as the number of regions decreases. This is consistent with our previous hypothesis on  $C_{prp}$ . This cost function tends to fit the lowest data in priority. Larger regions suffer more from this bias because they are constrained by more data, likely to be more dispersed. The confidence intervals within 1 standard deviation based on  $C_{prp}$  fail to overlap. Only the logarithmic least-squares method produces very similar fluxes whatever the number of regions, with all confidence intervals overlapping. The global flux based on  $C_{log}$ again seems to be the most reliable.

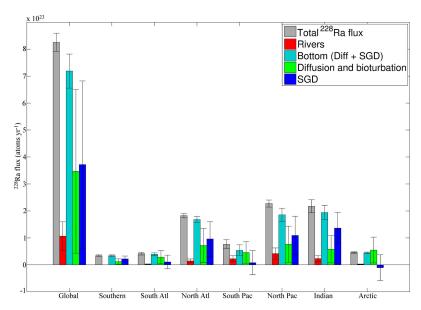
#### 3.4 Submarine groundwater discharge estimates

The shelf fluxes after inversion combine groundwater discharge, riverine particles, diffusion from sediments, and bioturbation. Here we deduce the contribution from groundwater discharge by using existing estimates of the other sources of radium.

Rivers are poor in dissolved <sup>228</sup>Ra and transport <sup>228</sup>Ra mainly with the sediments they carry (Moore and Shaw, 2008). According to Moore et al. (2008), the average dissolved  $^{228}$ Ra activity in rivers is  $0.65-1.95 \times 10^5$  atoms  $1^{-1}$ . As rivers annually discharge 35 000 km<sup>3</sup> of freshwater (Milliman, 2001), the dissolved riverine source lies between 2.3 and  $6.8 \times 10^{21}$  atoms yr<sup>-1</sup>, which is less than 1 % of the total flux. Nearly all the particulate radium is desorbed in the mixing zone, because of the salinity increase. According to various studies (Key et al., 1985; Moore et al., 1995; Krest et al., 1999), the amount of <sup>228</sup>Ra desorbed per gram of sediment lies in the range of 2.9 to  $8.7 \times 10^6$  atoms. In this study, we follow Milliman (2001) who proposed a global river sediment flux of  $1.8 \times 10^{16}$  g yr<sup>-1</sup>, divided into fluxes from six basins: South Atlantic, North Atlantic, South Pacific, North Pacific, Indian Ocean, and Arctic. This leads to a global  $^{228}$ Ra river flux estimate of  $0.53-1.60 \times 10^{23}$  atoms yr<sup>-1</sup>, significant but not dominant.

Radium-228 is released from the sediments by diffusion, bioturbation, and advection, the latter being associated with the SGD. Like Moore et al. (2008), we assume that the <sup>228</sup>Ra fluxes by diffusion and bioturbation from relict sands, composing 70 % of the total continental shelf area, is weak, typically of the order of  $10^9$  atoms m<sup>-2</sup> yr<sup>-1</sup> (Colbert, 2004; Hancock et al., 2006). Fluxes from continental shelf muds, which correspond to the remaining 30%, have been estimated by several studies using different methods (Table 6), such as inventories (Moore et al., 1995), benthic chambers (Hancock et al., 2000), sediment profiles (Hancock et al., 2000), or modeling (Hancock et al., 2006). But some of them were done at a time when SGD were not considered to be an important source of radium and may have included them in their estimates. These estimates should then be considered as an upper limit. Furthermore, their locations are often very close to the coast. Some recent studies, addressing diffusive fluxes (from muds or broader regions containing both types) separately, in order to estimate local groundwater discharge (Crotwell and Moore, 2003; Kim et al., 2005), used the simplified equation of diffusion from Krest et al. (1999) (requiring only water diffusion, sediment porosity, decay constants, and production rates) and pointed out to significantly lower values, but may not be representative of all continental shelves. So far, it is not possible to be very precise and a wide range has to be considered: from the low recent values, close to  $5 \times 10^9$  atoms m<sup>-2</sup> yr<sup>-1</sup>, to the higher values typically ranging from 25 to  $75 \times 10^9$  atoms m<sup>-2</sup> yr<sup>-1</sup> (Moore et al., 2008). The full range is then  $5-75 \times 10^9$  atoms m<sup>-2</sup> yr<sup>-1</sup>. As the continental shelf area in our model is  $2.73 \times 10^{13}$  m<sup>2</sup>, this means a total flux due to diffusion and bioturbation ranging from 0.43 to  $6.51 \times 10^{23}$  atoms yr $^{-1}$ . Then, using the logarithmic cost function, the SGD  $^{228}$ Ra flux estimate varies between 0.62 and  $6.82 \times 10^{23}$  atoms yr<sup>-1</sup>: these two fluxes are within the same range.

Figure 8 shows the <sup>228</sup>Ra fluxes for seven basins corresponding to the Southern Ocean and the six basins used by Milliman (2001) to define sediment inputs by rivers. The largest <sup>228</sup>Ra fluxes are found in the North Atlantic, the North Pacific, and the Indian Ocean, in roughly equal proportions, whereas those from the South Atlantic, the South Pacific, and the Arctic Ocean are far smaller. Fluxes from the last three



**Figure 8.** <sup>228</sup>Ra sources by basin, divided into riverine input, diffusion from sediment and SGD. Only logarithmic least-squares results are shown. Fluxes are divided the following way: Regions 1 and 38: Southern; 2–4: South Atl; 5–16: North Atl; 17–18 and 26–27: South Pac; 19–25: North Pac; 28–34: Indian; 35–37: Arctic.

**Table 6.** Estimates of  $^{228}$ Ra diffusive fluxes ( $10^9$  atoms m $^{-2}$  yr $^{-1}$ ) from fine-grained sediments (muds) by different methods. The studies which do not separate SGD and diffusion are in bold.

Reference	Study area	Method	Flux
Santschi et al. (1979)	Narragansett Bay	Inventory	52
Moore et al. (1995)	Amazon shelf	Inventory	110
Rama and Moore (1996)	North Inlet	Sediment profile	5
Hancock et al. (2000)	Zecks Lagoon	Benthic chamber Sediment profile	80 80
Crotwell and Moore (2003)	Port Royal Sound Marsh	Krest formula	9
Hancock et al. (2006)	Greet Barrier Reef	Modeling	26

regions are low enough in fact to be within the confidence intervals of riverine and diffusive sediment <sup>228</sup>Ra fluxes, so that the SGD confidence intervals include negative values, which are physically impossible.

SGD  $^{228}$ Ra end-member concentration varies considerably from one aquifer to another, ranging from 0.04 to  $125 \times 10^6$  atoms m<sup>-3</sup> (Moore et al., 2008). Measurements so far have shown aquifer concentrations in the Atlantic Ocean higher than the average by typically 30% (Kwon et al., 2014). Following Kwon et al. (2014), we assume that the aquifer concentrations are log-normally distributed with an average of 0.98 to  $1.15 \times 10^3$  dpm m<sup>3</sup>. Taking a geometric mean rather than an arithmetic mean is implicitly assuming that aquifers characterized by high  $^{228}$ Ra concentrations are emitting less water. It also suffers from wells being concentrated in developed countries. Using the results from the in-

version technique based on the logarithmic cost function, the total SGD flux is estimated to  $1.3{\text -}14.7 \times 10^{13} \, \text{m}^3 \, \text{yr}^{-1}$ , to be compared with the global river flow of  $3.5 \times 10^{13} \, \text{m}^3 \, \text{yr}^{-1}$  (Milliman, 2001). As long as diffusion cannot be separated from SGD more efficiently and is not shown to be negligible, no precise SGD estimate based on  $^{228}$ Ra can be produced. According to Taniguchi et al. (2007), there is only  $2.6 \times 10^{12} \, \text{m}^3 \, \text{yr}^{-1}$  of fresh groundwater discharge, representing between 2 and 20 % of the SGD, and the remaining part is recirculated seawater.

In total, 63 % of the  $^{228}$ Ra flows into the Indian and Pacific basins, which is more than their share of the continental shelf (52 %). As shown in Fig. 8, the estimated contributions of SGD to  $^{228}$ Ra fluxes are in very similar proportions, with 68 % going to the Indian and Pacific basins. This represents more than 70 % of all the groundwater discharge when the

differences in groundwater <sup>228</sup>Ra concentration are considered. However, uncertainties on SGD distribution are high, because diffusion from sediment might not be identical in all basins. These results contrast with the other source of water, nutrients and trace elements to the ocean: 60 % of the global river discharge flows to the Atlantic and Arctic basins (Milliman, 2001).

#### 4 Discussion

#### 4.1 Comparisons with previous studies

Our inversion estimates are in good agreement with previous regional studies of <sup>228</sup>Ra based on inventories. The <sup>228</sup>Ra inventory of the Mediterranean Sea computed by Rodellas et al. (2015) led to a total flux due to rivers, sediments and groundwater discharge of  $1.86-2.48 \times 10^{22}$  atoms yr<sup>-1</sup>. This is compatible with our estimate based on  $C_{log}$ , 1.96–  $2.28 \times 10^{22}$  atoms yr<sup>-1</sup>. Other cost functions produce an underestimation in the model by a factor 2.5 (ols) or 3.8 (prp). In the Yellow Sea, Kim et al. (2005) estimated a flux of  $3.3 \times 10^{15} \,\mathrm{dpm} \,\mathrm{yr}^{-1}$ , or  $1.4 \times 10^{22} \,\mathrm{atoms} \,\mathrm{yr}^{-1}$ . Expressed per unit of surface, it corresponds to  $3.6 \times 10^{10}$  atoms m<sup>-2</sup> vr<sup>-1</sup>. This is slightly lower than our flux for the larger region 23 (Sea of Japan, Yellow Sea, East China Sea) of 4.2- $5.6 \times 10^{10}$  atoms m<sup>-2</sup> yr<sup>-1</sup> based on  $C_{log}$ , both these fluxes being larger than the global average. At larger scale, Moore et al. (2008) estimated the total <sup>228</sup>Ra flux over the Atlantic between  $50^{\circ}$  S and  $80^{\circ}$  N to be  $2.8-4.2\times10^{23}$  atoms yr<sup>-1</sup>. Restricting the inversion to the Atlantic with  $C_{\log}$  yields 2.64–  $2.92 \times 10^{23}$  atoms yr<sup>-1</sup>, in the lower range but compatible. At least two reasons can explain the difference. Moore et al. (2008) have included data down to 1000 m, which allowed them to estimate <sup>228</sup>Ra release from sediments down to that depth. The authors have estimated the sources between 200 and  $1000 \,\mathrm{m}$  deep to  $0.13-0.37 \times 10^{22} \,\mathrm{atoms} \,\mathrm{yr}^{-1}$ . Furthermore, in order to compute the total <sup>228</sup>Ra content of the Atlantic, they have performed a linear interpolation of the data, potentially leading to errors, especially in areas where measurements are sparse.

At the global scale, Kwon et al. (2014) have used a method which is quite similar to ours and they have computed a total global  $^{228}\text{Ra}$  flux of  $9.1\text{--}10.1\times10^{23}$  atoms  $yr^{-1}$ , of which between 4.2 and  $7\times10^{23}$  atoms  $yr^{-1}$  are released by SGD. This corresponds to a global SGD flux of  $9\text{--}15\times10^{13}~\text{m}^3~\text{yr}^{-1}$ . The flux of  $^{228}\text{Ra}$  estimated by the present study is thus significantly lower than the estimates by Kwon et al. (2014), although our results include the Arctic and Antarctic sources (0.43–0.50 and 0.28–0.35 $\times10^{23}$  atoms  $yr^{-1}$  respectively). Kwon et al. (2014) minimized an ordinary least-squares cost function with 50 regions. We have shown here that both a high number of source regions and the use of the ordinary least-squares cost function concur to produce a higher estimate. However, this is at the expense of a higher uncer-

tainty and of producing unrealistic negative fluxes in some source regions. Additional differences in Kwon's study may explain their higher estimates such as a different ocean circulation model, a coarser vertical resolution, and a bathymetry re-gridded onto the model domain. Finally, they added dust deposition and removed data higher than 140 dpm m<sup>-3</sup>, but these two factors should rather tend to reduce their fluxes.

As recently proposed by Charette et al. (2016), shelf <sup>228</sup>Ra fluxes can be used as gauges of shelf fluxes of trace elements and isotopes, including nutrients, iron, and rare earth elements. <sup>228</sup>Ra is particularly relevant because it is chemically conservative and integrates information over annual to decadal timescales. At first approximation, the flux of TEIs is deduced from the <sup>228</sup>Ra flux and the ratio of the nearshore gradients. Limited for now due to the lack of nearshore measurements, this method could be more common in the future. As they are based on realistic assumptions on error statistics and have low uncertainties, our radium fluxes are able to improve the current estimates of all elements originating from the continental shelf.

In contrast, our uncertainties on groundwater discharge are large, even when compared to previous estimates. These larger uncertainties stem from the poor knowledge of the non-SGD sources of radium that we subtract from the total flux. As diffusion and bioturbation are expressed in flux per area, the mean SGD fluxes and their uncertainties depend to a large extent on the radium-emitting area we consider. Based on a more realistic refined bathymetry than Kwon et al. (2014)  $(2.73 \times 10^{13} \text{ m}^2 \text{ compared to } 1.5 \times 10^{13} \text{ m}^2)$ , our study also has a larger sedimentary flux, with an upper range close to the total <sup>228</sup>Ra flux. Thus, the lower range of SGD fluxes,  $1.3 \times 10^{13}$  m<sup>3</sup> yr<sup>-1</sup>, is very low, while the upper range,  $14.7 \times 10^{13} \,\mathrm{m}^3 \,\mathrm{yr}^{-1}$ , is similar to other studies. The use of this model provides an upper estimate but cannot precisely compute the global SGD flux, and no inverse model can if the surface diffusive flux and the area emitting radium are not clarified. Our expectation is that the lower range of sedimentary flux is more likely, because it comes from studies where SGD is also considered and because it never produces negative regional fluxes. By contrast, the higher range is very similar to the total bottom flux, as expected when no distinction is made between diffusion and SGD.

Comparisons with local direct estimates of SGD based on seepage meters and piezometers are also possible but less conclusive because of the high spatial variability of SGD. Our global average SGD flux lies between 0.5 and 5.5 m yr<sup>-1</sup>. Values from seepage meters and piezometers in the upper 200 m reported by Taniguchi et al. (2002) and Knee and Paytan (2011) range from 0.03 m yr<sup>-1</sup> in the Tokyo Bay (Taniguchi et al., 2002) to 1790 m yr<sup>-1</sup> near Mauritius (Burnett et al., 2006). Most measurements have been performed in shallow coastal areas, less than 10 m deep, and range from 1 to 50 m yr<sup>-1</sup>, 1 order of magnitude higher than our averages on all the seafloor down to 200 m deep. Thus, this range is very wide and many local studies span several orders of mag-

nitude. At a global scale, Taniguchi et al. (2009) produced an estimate of the global SGD flux,  $6.1-12.8 \times 10^{13} \, \mathrm{m^3 \, yr^{-1}}$ , in the upper range of our global estimate, by extrapolating seepage measurements. However, it suffers from most measurements being concentrated in developed countries.

The spatial distribution of the fluxes in this study is consistent with Kwon et al. (2014), with two-thirds of radium-228 flowing to the Indian and Pacific basins and also quite high fluxes in the western North Atlantic. It turns out that these basins have the highest riverine sediment loads (Milliman, 2001). However, we have shown that rivers could account for 7 to 20 % of <sup>228</sup>Ra fluxes only, which does not exclude an indirect impact on the other sources through the geology of the continental shelf. A difference among basins in the quantity of radium diffusion, which is poorly known, is also possible. If it is not the case, then the SGD are significantly higher in East Asia than in other regions. SGD can be driven by as many factors as storms, waves, tides, and thermal gradients, and depend on the permeability and structure of coastal and shelf sediments (Moore, 2010a). Comparative data on these factors in several basins will be necessary to explain the origin of the observed differences.

#### 4.2 Model biases

Our model of <sup>228</sup>Ra is based on a circulation model and assumptions on the cycle of this isotope. Both are potential sources of errors.

Dust has not been included in the model. The model then replaces them with other sources, potentially leading to an overestimation. At global scale, it represents  $1.7 \times$  $10^{21} \, \text{atoms yr}^{-1}$  (Kwon et al., 2014), less than 0.2 % of the total flux and less than most individual source regions, which cannot create large biases. Nevertheless, the largest dust deposition is associated to Saharan dust transported to the North Atlantic, in the Canary upwelling region (Mahowald et al., 2005). In this region, dust may have some impact on the <sup>228</sup>Ra distribution, since it brings this isotope directly to the open ocean. Its absence in our study might explain the overestimation by our model near the coast of region 6 and underestimation offshore. Yet, the exclusion of dust deposition in our analysis cannot explain the largest bias in the North Atlantic: the overestimation in the Gulf Stream coupled with an underestimation just southeast (see Sect. 3.2), because the area where the model cannot transport radium is located west of the maximum of dust input and displays higher concentration.

Scavenging is a neglected potential sink in this study. As the residence time related to scavenging is approximately 500 years (Moore and Dymond, 1991), it accounts for between 1 and 2% of all sinks. Thus, its inclusion in our computations would increase the source intensity necessary to maintain global balance. Contrary to radioactivity, scavenging is highly heterogeneous, most intense where primary productivity and particle concentrations are highest. Fluxes from

the high latitudes of the North Atlantic are thus potentially more underestimated than fluxes in other regions, since they are areas of intense biological activity during spring blooms. As the actual total lifetime of <sup>228</sup>Ra is overestimated when scavenging is not taken into account, the gradient between coast and open ocean could be too low. However, as the horizontal mixing timescale of the ocean is a few decades, the relative overestimation of open ocean concentrations is less than 10 %.

The contribution of rivers to radium fluxes is considered when estimating the SGD, but only at a basin-wide scale. As most riverine <sup>228</sup>Ra travels attached to particles and is desorbed in the mixing zone, we have based our computation on the sediment loads of Milliman (2001), which are basin-wide. Although NEMO 3.6 takes rivers into account for their impact on salinity (Madec, 2015), for now no information at the model grid scale is available on their sediment loads. If this information existed, we would have been able to estimate the contribution of rivers, and consequently sediments and SGD, in the <sup>228</sup>Ra fluxes for each of the 38 regions. Some local SGD fluxes close to large rivers (Amazon, Congo, Yellow River, etc.) could then appear to be significantly lower than their shares of the total <sup>228</sup>Ra fluxes suggest.

The other part of the model is the circulation model. The climatological circulation of NEMO 3.6 was not optimized in this study. However, the residuals after inversion show that some regions are associated with spatially structured residuals. There are good reasons to incriminate the ocean circulation. Because of the low resolution of the model (2°), isopycnal diffusion has been used to parameterize sub-grid processes and mixing (Redi, 1982) with a constant eddy diffusivity of 2000 m<sup>2</sup> s<sup>-1</sup>. Nevertheless, in very energetic regions, such as the western boundary currents (WBCs; e.g., the Gulf Stream and the Kuroshio), higher eddy diffusivity might be needed to enhance the exchanges with the open ocean, possibly improving the fit with the observations. Furthermore, it is well known that the mean currents are also dependent on this low resolution (impacting for instance the position and intensity of the WBCs). Although it cannot solve all the flaws of the circulation, improvements could thus be brought by an increase of the resolution towards eddy resolving models. However, as 100 years of simulation are needed at a global scale for each source region, this would be computationally expensive.

Other sources of errors are the four statistical assumptions on the errors: errors are assumed to have zero expectation, no correlation, a normal distribution, and variances depending on the concentrations in a way specific to each cost function. Systematic biases on data or model are not corrected by least-squares algorithms. They increase or decrease values without leaving clues. However, the model conserves mass: the quantity of radium present in the ocean from each model tracer is precisely known and if concentration is too high at some place, it will be too low elsewhere. As for the measurements, their uncertainty is generally around 10% or lower,

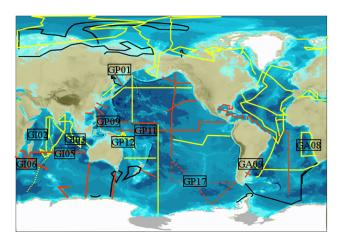
and cruises take them independently from each other, making the assumption of zero expectation on the observational errors reasonable. The second assumption is the absence of prior correlation. If prior uncertainties of neighboring data are correlated, it means that the errors are likely to be of the same sign whatever the solution, and that multiplying measurements in this area does not multiply information proportionally. Where measurements are dense, with residuals far from the expected white noise, for instance in the North Atlantic, there may be correlation and uncertainties may be underestimated. The last assumptions are about the structure of variance. We have shown that logarithmic residuals were almost normally distributed and independent from concentrations (see Sect. 3.2), justifying the choice of  $C_{log}$  as a cost function. Cols leads to higher uncertainties, especially for small sources, and  $C_{prp}$  to systematic underestimation, but both are useful for comparison, in order to identify regions where physical assumptions are inaccurate (see Sect. 3.1).

#### 4.3 What new data would be most useful?

Observations are not evenly distributed. Some coastal regions cannot be constrained properly because <sup>228</sup>Ra data are lacking. Improving the coverage would increase the quality of the inversion in two ways: it would reduce the uncertainties and make it possible to divide wide regions into smaller regions as long as they have distinct footprints. For instance, the Philippines, Papua, or the Gulf of Guinea, whose footprints are very different from the Indonesian seas and the southwestern coast of Africa, could be considered as independent regions. More samples in the Indian Ocean, the South Atlantic (south of 30°S), the Southern Ocean, and the western equatorial Pacific are priorities. All these regions will be sampled by upcoming GEOTRACES cruises shown in Fig. 9. At the same time, deep samples will be taken outside of the Atlantic, enabling a more comprehensive global inversion with extra source regions at greater depths. This would improve our knowledge of the contribution of deeper sediments.

Most direct submarine groundwater discharge measurements have been performed in developed countries, with a focus on the North Atlantic and the Mediterranean Sea (Taniguchi et al., 2002). At the same time, measurements are completely lacking over large regions. More SGD studies in areas where they are potentially the highest, namely the Bay of Bengal, the Indonesian seas and the China seas, would produce more representative estimates of the <sup>228</sup>Ra content of SGD around the world and direct estimates of local SGD magnitude to be compared with regional inversion results. They have begun more recently (Knee and Paytan, 2011) and are still sparse.

Information contained in <sup>228</sup>Ra might be completed with <sup>226</sup>Ra concentrations, measured during the same campaigns and for this reason available with a similar coverage. Associated with the same source as <sup>228</sup>Ra, but with a much longer



**Figure 9.** Map of GEOTRACES cruises (from http://www.geotraces.org/cruises/cruise-summary). Planned sections are in red, completed sections in yellow, International Polar Year sections in black. Names in rectangles correspond to cruises potentially bringing the largest extra information on <sup>228</sup>Ra fluxes because they are performed in places lacking measurements.

half-life, 1602 year, <sup>226</sup>Ra would constrain deeper sources and would help in assessing the quality of the thermohaline circulation and deep ventilation of the circulation model. However, inverting <sup>226</sup>Ra data would require a precise modeling of scavenging, which is not negligible at these longer timescales, as well as a much longer integration duration in order to get steady state distribution.

# 5 Conclusions

Based on inverse modeling, we have computed a global  $^{228}$ Ra flux from continental shelves of  $8.01-8.49 \times$  $10^{23}$  atoms yr<sup>-1</sup>, with the largest sources in the western Pacific, the western North Atlantic, and the Indian oceans and the smallest sources in the eastern Pacific. The Arctic and Antarctic sources have been estimated for the first time, accounting for 0.43–0.50 and 0.28–0.35  $\times$  10<sup>23</sup> atoms yr<sup>-1</sup> respectively. These precise estimates are obtained by minimizing the squared differences between model and observed concentrations on a logarithmic scale. We think this cost function is more realistic than a linear least-squares one because it assumes that the error standard deviation is proportional to <sup>228</sup>Ra concentration, rather than constant, which proves to be more accurate. It is also better than a "proportional" cost function, weighted by the inverse of the observed concentration, which tends to produce underestimations of concentrations and fluxes. Given the number of available measurements, we were able to constrain 38 regional fluxes. The shelf fluxes produced using these optimal parameters are around 20 % lower than previous estimates. In a near future, they will enable us to quantify continental shelf fluxes of trace elements and isotopes to the oceans at any place where nearshore gradients are measured (Charette et al., 2016). The estimated global SGD is far less precise, ranging between 1.3 and  $14.7 \times 10^{13} \,\mathrm{m}^3 \,\mathrm{yr}^{-1}$ , because of the very large uncertainty on the two other sources of <sup>228</sup>Ra - i.e., riverine particles and most of all diffusion from bottom sediments, also located on the continental shelf. Only the upper range is compatible with previous estimates. These results mean that all the studies which have used or will use <sup>228</sup>Ra to estimate SGD will suffer from a very high uncertainty until diffusion is properly estimated or proven to be negligible. After inversion, we were able to reproduce the basin-scale patterns of <sup>228</sup>Ra distribution with nevertheless systematic biases in several regions, especially in the Arctic, and west of the subtropical gyres. Shortcomings in the circulation model are the most probable explanation of these biases (too weak exchanges between continental shelves and open ocean). Extensive regions are lacking observations, mostly in the Southern Hemisphere (Pacific, Indian, and mostly Southern Ocean, as well as western equatorial Pacific Ocean), and better coverage in basins where SGDs are known to be influential and to produce large horizontal gradients is also needed (such as in South Asia). But the main impediment to achieve precise estimates of global Ra SGD fluxes comes from the very poor knowledge of diffusive sedimentary fluxes: without a proper way to separate diffusion and SGD, inversions can compute the total bottom flux but are not able to precisely evaluate these two components.

Code and data availability. The source code of NEMO is available on the NEMO website (http://www.nemo-ocean.eu). Studies providing data are listed in Table S1 in the Supplement. The inversion code and data used in this study can be obtained directly by contacting the authors.

## Appendix A: Error statistics

This section describes the way regional  $^{228}$ Ra flux estimates and their error bars are computed. The inverse problem is the following one: using p measurements of oceanic  $^{228}$ Ra concentrations, their n sources must be traced back, by means of a circulation model. The model produces n  $^{228}$ Ra concentration fields, one from each of the source regions. It is run through steady state, the ith field reaching concentration  $A_i$ . The model concentration is a linear combination of the latter:

$$[^{228}\text{Ra}]^{\text{mod}} = \sum_{i=1}^{n} A_i x_i.$$

Inversion requires fitting the model concentration to the observations. As the expression of model concentration is linear in terms of the unknowns, the problem can be written

$$\mathbf{A}\mathbf{x} = \mathbf{b} + \mathbf{\eta}$$

where  $b \in \mathcal{M}_{p,1}$  is the vector of the p measured radium concentrations,  $x \in \mathcal{M}_{n,1}$  is the unknown and the vector of the n source intensities,  $\mathbf{A} \in \mathcal{M}_{p,n}$  is the footprint matrix, composed of the  $A_i$  at each data point, and  $\eta \in \mathcal{M}_{p,1}$  is the vector of residuals, accounting for the misfits, which are inevitable in our over-constrained problem (p > n).

We want to minimize the distance between data and model concentrations, summed up in a cost function. The ordinary least-squares cost function we first used is given by

$$C_{\text{ols}}(\mathbf{x}) = (\mathbf{A}\mathbf{x} - \mathbf{b})^{\top} (\mathbf{A}\mathbf{x} - \mathbf{b}).$$

The cost function has to be minimal at the optimal x, called  $x_{\text{opt}}$ , implying, in the absence of inequality constraints,

$$\nabla C_{\text{ols}}(\mathbf{x}_{\text{opt}}) = 2\mathbf{A}^{\top}(\mathbf{A}\mathbf{x} - \mathbf{b}) = \mathbf{0}.$$

As the  $n^{228}$ Ra concentration fields are independent from each other,  $\mathbf{A}^{\mathsf{T}}\mathbf{A}$  is invertible. Then,

$$\mathbf{x}_{\text{opt}} = \mathbf{A}_{\text{inv}} \mathbf{b}$$

$$\mathbf{A}_{\text{inv}} = (\mathbf{A}^{\top} \mathbf{A})^{-1} \mathbf{A}^{\top}.$$
(A1)

 ${\bf A}_{\rm inv}$  is the pseudo-inverse of  ${\bf A}$ , transforming the consequences into their most probable causes and justifying the word "inversion". Equation (A1) corresponds to (2.95) in Wunsch (2006). In practice, fluxes are constrained to be positive. This is managed by just reducing the number of source regions if necessary, without changing the principle of the inversion.

Matrix H is defined by

$$\mathbf{H} = \mathbf{A}(\mathbf{A}^{\top}\mathbf{A})^{-1}\mathbf{A}^{\top} = \mathbf{A}\mathbf{A}_{inv}.$$

As  $\mathbf{H}^2 = \mathbf{H}^\top = \mathbf{H}$ ,  $\mathbf{H}$  is an orthogonal projector on a subspace of dimension n. As  $\mathbf{A}\mathbf{x}_{\text{opt}} = \mathbf{H}\mathbf{b}$ , the model concentration is the projection of the observations on this subspace

whereas the vector of residuals is the projection on the complementary subspace.

Simple algebraic transformations yield

$$\mathbf{H}\mathbf{A} = \mathbf{A}(\mathbf{A}^{\top}\mathbf{A})^{-1}\mathbf{A}^{\top}\mathbf{A} = \mathbf{A}$$

$$\mathbf{A}_{\text{inv}}\mathbf{A} = (\mathbf{A}^{\top}\mathbf{A})^{-1}\mathbf{A}^{\top}\mathbf{A} = \mathbf{I}_{n}$$

$$\mathbf{A}_{\text{inv}}\mathbf{A}_{\text{inv}}^{\top} = (\mathbf{A}^{\top}\mathbf{A})^{-1}\mathbf{A}^{\top}\mathbf{A}(\mathbf{A}^{\top}\mathbf{A})^{-1} = (\mathbf{A}^{\top}\mathbf{A})^{-1}.$$
(A2)

Uncertainties on x resulting from uncertainties on data (b) and model (A) must now be evaluated. By principle of the ordinary least-squares cost function, the error covariance of (Ax - b) is considered to be a constant diagonal matrix:

$$\langle (\mathbf{A}\mathbf{x} - \mathbf{b})(\mathbf{A}\mathbf{x} - \mathbf{b})^{\top} \rangle = \sigma^2 \mathbf{I}_p.$$

The posterior covariance matrix of x is

$$\mathbf{V}_{xx} = \langle (\mathbf{x} - \mathbf{x}_{\text{opt}})(\mathbf{x} - \mathbf{x}_{\text{opt}})^{\top} \rangle$$

$$= \langle (\mathbf{A}_{\text{inv}} \mathbf{A} \mathbf{x} - \mathbf{A}_{\text{inv}} \mathbf{b}) (\mathbf{A}_{\text{inv}} \mathbf{A} \mathbf{x} - \mathbf{A}_{\text{inv}} \mathbf{b})^{\top} \rangle$$

$$= \mathbf{A}_{\text{inv}} \langle (\mathbf{A} \mathbf{x} - \mathbf{b}) (\mathbf{A} \mathbf{x} - \mathbf{b})^{\top} \rangle \mathbf{A}_{\text{inv}}^{\top}$$

$$= \sigma^{2} \mathbf{A}_{\text{inv}} \mathbf{A}_{\text{inv}}^{\top}.$$
(A3)

Equation (A3) corresponds to (2.102) in Wunsch (2006).  $\sigma^2$  is related to the root mean square of residuals  $v_{\eta}$  the following way:

$$v_{\eta} = \langle (\mathbf{A}\boldsymbol{x}_{\text{opt}} - \boldsymbol{b})^{\top} (\mathbf{A}\boldsymbol{x}_{\text{opt}} - \boldsymbol{b}) \rangle$$

$$= \langle (\mathbf{H}\boldsymbol{b} - \boldsymbol{b})^{\top} (\mathbf{H}\boldsymbol{b} - \boldsymbol{b}) \rangle$$

$$= \langle ((\mathbf{I}_{p} - \mathbf{H})(\mathbf{A}\boldsymbol{x} - \boldsymbol{b}) + (\mathbf{H}\mathbf{A} - \mathbf{A})\boldsymbol{x})^{\top} ((\mathbf{I}_{p} - \mathbf{H})(\mathbf{A}\boldsymbol{x} - \boldsymbol{b}) + (\mathbf{H}\mathbf{A} - \mathbf{A})\boldsymbol{x}) \rangle$$

$$= \langle (\mathbf{A}\boldsymbol{x} - \boldsymbol{b})^{\top} (\mathbf{I}_{p} - \mathbf{H})(\mathbf{A}\boldsymbol{x} - \boldsymbol{b}) \rangle$$

$$= \sigma^{2} \frac{p - n}{p}.$$
(A4)

Then, using Eqs. (A2), (A3) and (A4), we can compute  $V_{xx}$ :

$$\mathbf{V}_{xx} = \frac{p}{p-n} v_{\eta} (\mathbf{A}^{\top} \mathbf{A})^{-1}. \tag{A5}$$

The diagonal terms of  $V_{xx}$  are the squared uncertainties on x whereas the other terms are covariances, indicating source regions with similar footprints.

The result can be extended to cases where the cost function is weighted, such as the proportional least-squares inversion, by just normalizing matrix  $\mathbf{A}$  and vector  $\mathbf{b}$ . Extension to logarithmic inversion is done by linearizing  $\log(\mathbf{A}\mathbf{x}) - \log(\mathbf{b})$  to  $\mathbf{F}\mathbf{x} - \boldsymbol{\beta}$  with  $\boldsymbol{\beta} = \log(\mathbf{b}) + \mathbf{F}\mathbf{x}_{\text{opt}} - \log(\mathbf{A}\mathbf{x}_{\text{opt}})$ , assuming a small error.

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