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Optimal Riemannian quantization with an application to air traffic analysis

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Abstract

The goal of optimal quantization is to find the best approximation of a probability distribution by a discrete measure with finite support. In this paper, we introduce an algorithm that computes the optimal finite approximation when the probability distribution is defined over a complete Riemannian manifold. It is a natural extension of the well-known Competitive Learning Vector Quantization algorithm. We use it to analyze air traffic complexity in air traffic management (ATM). From a given air traffic situation, we extract an image of covariance matrices and use CLRQ (Competitive Learning Riemannian Quantization) to find the best finite approximation of their empirical spatial distribution. This yields a digest of the traffic as well as a clustering of the airspace into zones that are homogeneous with respect to complexity. These digests can then be compared using discrete optimal transport and be further used as inputs of a machine learning algorithm or as indexes in a traffic database.

1 Introduction

1.1 Motivation

Most of the statistical tools developed so far are dedicated to data belonging to vector spaces, since it is the most convenient setting for algorithms as well as theoretical derivations. However, when dealing with real world applications, such a framework may not fit in with the structure of the data. It is obviously the case for geostatistics over world-sized datasets, but it is also true in many other fields: shapes in computer vision, diffusion tensor images in neuroimaging, signals in radar processing do not belong to a Euclidean space, but rather to a differentiable manifold. Riemannian geometry provides a convenient framework to deal with such objects. It allows a straightforward generalization of the basic statistical notions such as barycenters (Karcher 1977; Fletcher et al. 2008), covariance and Gaussian distributions (Pennec 2006; Said et al. 2017), and of the usual linear operations such as principal component analysis (Fletcher et al. 2004; Sommer et al. 2010). The use of these statistical tools has met a growing interest in various fields, including shape analysis (Kendall 1984), computational anatomy (Fletcher et al. 2004), medical imaging (Fletcher and Joshi 2007), probability theory (Bigot et al. 2017), and Radar signal processing (Arnaudon et al. 2013; Le Brigant 2016).

In air traffic management (ATM), a major concern is the ability to infer an estimation of the perceived complexity for a controller from the knowledge of aircraft trajectories in a given airspace, as depicted in Figure 1. Many interdependent factors are involved in the cognitive process of a human controller, making the problem extremely difficult to solve, if even possible. However, there is a consensus among the experts on the importance of traffic disorder. In the present work, we assume

the spatial distribution of the aircraft velocities to be Gaussian and use the covariance function as an indicator for local complexity, in compliance with the above remark. This model leads us to consider mappings from the plane to the space of symmetric, positive definite (SPD) matrices, that will be termed as images (of SPD matrices) in the sequel. While the space of SPD matrices does not form a vector space, it is well known that it can be equipped with the structure of a Riemannian manifold (Pennec et al. 2006). Working directly with the images is an extremely computationally expensive task, that is unrealistic in practice. Moreover, the information provided is highly redundant, making it cumbersome to use in statistical analysis. To cope with this problem, a digest of the images will be produced using optimal quantization. The main idea is to model an image as a realization of a random field with values in the space of SPD matrices. The values collected at each sample point in the image will thus be representative of an underlying probability distribution, supported on the Riemannian manifold of the SPD matrices. Using optimal quantization on it will then allow a digest (or summary) to be produced, that can be seen as a generalization of the intensity histogram in classical image processing.

1.2 Optimal quantization

Indeed, optimal quantization is concerned with finding the best approximation, in the sense of the Wasserstein metric, of a probability distribution by a discrete measure with finite support. Its origins go back to information theory and signal processing in the middle of the twentieth century. It was initially developed for signal compression purposes, in order to find appropriate ways to discretize a signal. It has since been applied to many different problems, in various fields : in numerical integration, to approximate integrals with respect to a certain probability distribution μ by a finite sum; in finance, for the numerical computation of conditional expectations (Pagès et al. 2004); and in economics, to determine the optimal location of service centers with respect to population demand (Bouchitté et al. 2011). Furthermore, optimal quantization naturally yields a clustering, which coincides with the solution given by the K -means algorithm when μ is an empirical law equally distributed on a finite number of points. In most of these applications, quantization is performed on distributions of real d -dimensional random vectors; however, the quantization framework has also been extended to distributions of stochastic processes, for applications in finance (Pages 2008) or to perform clustering of random functional data such as spectra or speech recordings (Biau et al. 2008). For further references, we refer the reader to the monograph (Graf and Luschgy 2007) or the survey paper (Pagès 2015).

1.3 The non-Euclidean setting

Here we are interested in quantizing the distribution of SPD matrices. These matrices are obtained from an air traffic image such as those shown in Figure 1, by evaluating in each point of the image the covariance function of the velocity field. These covariance matrices reflect the distribution of the directions as well as the amplitude of the velocities, coded increasingly from yellow to red in Figure 1. Quantizing their distribution yields two desirable results : (1) a clustering of the airspace into different zones homogeneous with respect to complexity, and (2) a summary of the image through the optimal discrete approximation. The first one means that each point of the image is assigned a level of complexity according to the class it belongs to, and the second one, that different traffic images can easily be compared through the comparison of their summaries, e.g. using optimal transport. This is an interesting prospect, since it allows for the indexation of air traffic situations in view of creating automatized decision-making tools to help air traffic controllers.

As previously mentioned, the space of SPD matrices is not a Euclidean space but a Riemannian

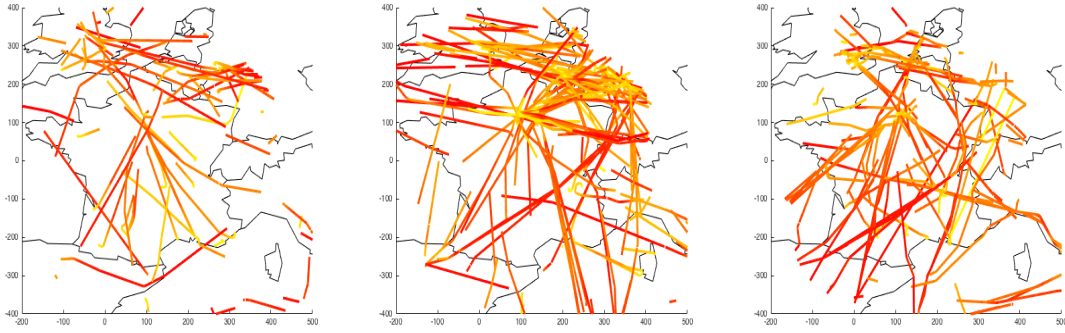


Figure 1: Traffic over the French airspace during one-hour periods of time.

nian manifold. Optimal quantization of probability distributions on Riemannian manifolds has recently received attention (Kloeckner 2012; Iacobelli 2016). In particular, the famous theorem on the asymptotic behavior of the quantization error, Zador’s theorem (see Graf and Luschgy 2007, Theorem 6.2), was extended to probability measures on a complete Riemannian manifold in (Iacobelli 2016). However, to the best of our knowledge, no numerical schemes have yet been introduced to find the optimal approximation of manifold-valued data by a smaller number of points. Here, we introduce Competitive Learning Riemannian Quantization (CLRQ), an algorithm that computes the optimal discrete approximation of a probability measure μ on a manifold, using on-line observations sampled from μ . It is a natural generalization of Competitive Learning Vector Quantization (CLVQ) (see e.g. Pagès 2015), a stochastic gradient descent algorithm exploiting the differentiability of the quantization cost function. The algorithm converges when the manifold is complete and μ is compactly supported. We apply it to distributions of covariance matrices extracted from real traffic datasets, and compare the results to those of vector quantization. We then compare different traffic situations by computing the optimal transport cost between their discrete summaries.

1.4 Outline of the paper

The paper is organized as follows. In Section 2, we briefly give some basic notions of Riemannian geometry, and introduce the context of optimal quantization. In Section 3, we present the CLRQ algorithm and give some theoretical results. After showing some simple illustrations on the circle, the sphere and the hyperbolic plane in Section 4, we present our main application in air traffic management in Section 5.

2 Mathematical setup

Let us begin by introducing some notations and reminding some basic notions of Riemannian geometry. We consider a differentiable manifold M of dimension d equipped with a Riemannian metric, i.e. a smoothly varying inner product $\langle \cdot, \cdot \rangle_x$ defined on each tangent space $T_x M$ at $x \in M$. Recall that $T_x M$ is a linear approximation of M at point x , and contains all tangent vectors to M at x . The subscript x in the metric will often be omitted and the norm associated to the Riemannian metric $\langle \cdot, \cdot \rangle$ will be denoted by $\| \cdot \|$. Letting ∇ be the Levi-Civita connection, the geodesics of M are the curves γ satisfying the relation $\nabla_{\dot{\gamma}} \dot{\gamma} = 0$, which implies that their speed has constant norm

$\|\dot{\gamma}(t)\| = \text{cst}$. They are also the local minimizers of the arc length functional l :

$$l: \gamma \mapsto \int_0^1 \|\dot{\gamma}(t)\| dt$$

where in the previous expression curves are assumed, without loss of generality, to be defined over the interval $[0, 1]$. The exponential map at point x is the mapping, denoted by \exp_x , that maps a tangent vector v of an open set $U \subset T_x M$ to the endpoint $\gamma(1)$ of the geodesic $\gamma: [0, 1] \rightarrow M$ starting at $\gamma(0) = x$ at speed $\dot{\gamma}(0) = v$. That is, $\exp_x(v) = \gamma(1)$. The inverse of the exponential map $\exp_x^{-1}(y)$, which gives the vector that maps x to y , will be denoted by the more intuitive notation $\vec{x}y$. We assume that M is complete, i.e. that the exponential map at x is defined on the whole tangent space $T_x M$. By the Hopf-Rinow theorem, we know that M is also geodesically complete, that is, any two points $x, y \in M$ can be joined by a geodesic of shortest length. This minimal length defines the geodesic distance between x and y , denote in the sequel by $d(x, y)$. For further details, we refer the reader to a standard textbook such as (Jost 2008).

Consider a probability measure μ on M with density and a compact support $K = \text{supp}\mu$, and let X be a random variable of law μ . We want to approximate X by a *quantized version* of X , i.e. a certain $q(X)$ where $q: M \rightarrow \Gamma \subset M$ is a measurable function with an image Γ of cardinal at most n . The function q is called an *n-quantizer* and is chosen to minimize the L^p criteria

$$\mathbb{E}_\mu [d(X, q(X))^p]. \quad (1)$$

It is well known that, since any n -quantizer q of image $\Gamma \subset M$ verifies for all $x \in M$, $d(x, q(x)) \geq \inf_{a \in \Gamma} d(x, a)$, with equality if and only if $q(x) = \arg \min_{a \in \Gamma} d(x, a)$, the optimal quantizer is the projection to the nearest neighbor of Γ . If $|\Gamma| < n$ and $|\text{supp}\mu| \geq n$, one easily checks that q can always be improved, in the sense of criteria (1), by adding an element to its image. Therefore, the optimal n -quantizer for criteria (1) is of the form $q_\Gamma: M \rightarrow \Gamma = \{a_1, \dots, a_n\}$, where the a_i 's are pairwise distinct, and

$$q_\Gamma(\cdot) = \sum_{i=1}^n a_i \mathbf{1}_{C_i(\Gamma)}(\cdot), \quad \text{where} \quad C_i(\Gamma) = \{x \in M, d(x, a_i) \leq d(x, a_j) \forall j \neq i\}.$$

The set $C_i(\Gamma)$ is the i^{th} Voronoi cell associated to Γ and the union of all these cells form the Voronoi diagram. The quantization problem is therefore equivalent to the approximation of the support of the probability measure μ by an n -tuple $(a_1, \dots, a_n) \in M^n$ minimizing the cost function $F_{n,p}: M^n \rightarrow \mathbb{R}_+$,

$$F_{n,p}(a_1, \dots, a_n) = \mathbb{E}_\mu \left(\min_{1 \leq i \leq n} d(X, a_i)^p \right) = \int_M \min_{1 \leq i \leq n} d(x, a_i)^p \mu(dx). \quad (2)$$

This cost function is called the *distorsion function* of order p for μ . Finally, there is a third way, in addition to (1) and (2), of expressing the quantization problem: it is also equivalent to the approximation of the measure μ by the closest discrete measure with respect to the Wasserstein distance of order p

$$W_p(\mu, \nu) = \inf_P \int d(u, v)^p dP(u, v). \quad (3)$$

Here ν is a discrete measure with a finite support of n points, and the infimum is taken over all measures P on $M \times M$ with marginals μ and ν . This is well known for the vector case (Graf and Luschgy 2007) and applies verbatim to measures on manifolds.

Since we have assumed that μ has compact support, the existence of a minimizer of $F_{n,p}$ is easy to obtain. For the sake of completeness, this is shown in the Appendix. The elements of a minimizer α are called *optimal n -centers* of μ . The minimizer α is in general not unique, first of all because any permutation of α is still a minimizer, and secondly because any symmetry of μ , if it exists, will transform α into another minimizer of $F_{n,p}$. For example, any rotation of the optimal n -centers of the uniform distribution on the sphere conserves optimality.

The n -th quantization error is defined by

$$V_{n,p}(\mu) = \inf_{(a_1, \dots, a_n) \in M^n} F_{n,p}(a_1, \dots, a_n) = \inf_{(a_1, \dots, a_n) \in M^n} \int_M \min_{1 \leq i \leq n} d(x, a_i)^p \mu(dx).$$

In the vector case, Zador's theorem (see Graf and Luschgy 2007, Theorem 6.2) tells us that it decreases to zero as $n^{-p/d}$. The limit of $n^{p/d} V_{n,p}(\mu)$ is proportional to the p^{th} quantization coefficient, i.e. the limit (which is also an infimum) when μ is the uniform distribution on the unit square of \mathbb{R}^d

$$Q_p([0, 1]^d) = \lim_{n \geq 1} n^{p/d} V_{n,p}(\mathcal{U}([0, 1]^d)).$$

Moreover, when μ is absolutely continuous with density h , the asymptotic distribution of the optimal discrete approximation is proportional to $h^{d/(d+p)}$.

In the case of a Riemannian manifold M , the moment condition of the flat case generalizes to a condition involving the curvature of M . The following term measures the maximal variation of the exponential map at $x \in M$ when restricted to a $(d-1)$ -dimensional sphere $S_\rho \subset T_x M$ of radius ρ

$$A_x(\rho) = \sup_{v \in S_\rho, w \in T_v S_\rho, \|w\| = \rho} \|d_v \exp_x(w)\|.$$

The following generalization of Zador's theorem to Riemannian quantization was proposed by Iacobelli (2016).

Theorem 1 (Iacobelli 2016, Theorem 1.4 and Corollary 1.5). *Let $(M, \langle \cdot, \cdot \rangle)$ be a complete Riemannian manifold without boundary, and let $\mu = h \, \text{dvol} + \mu_s$ be a probability measure on M , where dvol denotes the Riemannian volume form and μ_s the singular part of μ . Assume there exist $x_0 \in M$ and $\delta > 0$ such that*

$$\int_M d(x, x_0)^{p+\delta} d\mu(x) + \int_M A_{x_0}(d(x, x_0)^p) d\mu(x) < \infty.$$

Then

$$\lim_{n \rightarrow \infty} n^{p/d} V_{n,p}(\mu) = Q_r([0, 1]^d) \|h\|_{d/(d+p)},$$

where $\|\cdot\|_r$ denotes the L^r -norm. In addition, if $\mu_s = 0$ and (a_1, \dots, a_n) are optimal n -centers, then

$$\frac{1}{n} \sum_{i=1}^n \delta_{a_i} \xrightarrow{D} \lambda h^{d/(d+p)} dx \quad \text{as } n \rightarrow \infty,$$

where \xrightarrow{D} denotes convergence in distribution and λ is the appropriate normalizing constant.

In this work, we are interested in finding numerical schemes to compute the optimal n -centers $\alpha = (a_1, \dots, a_n)$ in practice from potentially large sets of data. To do so, we will search for the critical points of the distortion function.

3 Competitive Learning Riemannian Quantization

3.1 The algorithm

We assume that the only knowledge that we have of the probability measure μ that we want to approximate is through an on-line sequence of i.i.d. observations X_1, X_2, \dots sampled from μ . A classical algorithm used for quadratic ($p = 2$) vector quantization is the *Competitive Learning Vector Quantization* algorithm, a stochastic gradient descent method based on the differentiability of the distortion function $F_{n,2}$. We propose here a natural extension of this method to our setting, i.e. a compactly-supported probability measure on a complete Riemannian manifold. As shown below (Proposition 1), the distortion function is differentiable, and its gradient at point $\alpha = (a_1, \dots, a_n)$ is given by

$$\nabla_{\alpha} F_{n,2} = -2 \mathbb{E}_{\mu} \left(\mathbf{1}_{\tilde{C}_i}(X_k) \overrightarrow{a_i X_k} \right)_{1 \leq i \leq n},$$

with the notation $\overrightarrow{xy} = \exp_x^{-1}(y)$. In other words, the opposite direction of the gradient is, on average, given by the vectors inside the expectation. Competitive learning quantization consists in following this direction at each step k , that is, updating only the center a_i corresponding to the Voronoi cell of the new observation X_k , in the direction of that new observation. In the Riemannian setting, instead of moving along straight lines, we simply follow geodesics using the exponential map.

Assume that we have access to N independent and identically distributed observations X_1, \dots, X_N . We choose a sequence of positive steps $(\gamma_k)_{k \geq 1} \subset (0, 1)$ verifying the usual conditions

$$\sum_{k \geq 1} \gamma_k = +\infty, \quad \sum_{k \geq 1} \gamma_k^2 < +\infty. \quad (4)$$

We propose the following algorithm.

Algorithm 1 (Competitive Learning Riemannian Quantization).

Initialization : $\alpha(0) = (a_1(0), \dots, a_n(0))$.

For $k = 0, \dots, N - 1$,

1. find $i = \underset{j}{\operatorname{argmin}} d(X_{k+1}, a_j(k))$,
2. set $\alpha(k+1) = (a_1(k+1), \dots, a_n(k+1))$ where

$$\begin{aligned} a_i(k+1) &= \exp_{a_i(k)} \left(\gamma_{k+1} \overrightarrow{a_i X_{k+1}} \right), \\ a_j(k+1) &= a_j(k) \quad \forall j \neq i. \end{aligned}$$

The steps are chosen to be in $(0, 1)$ so that at each iteration, the center that is updated stays in the same Voronoi cell, guaranteeing that the centers stay pairwise distinct (if initially pairwise distinct). We will see in what follows that this algorithm converges, provided that the injectivity radius of M is lower-bounded by a positive constant.

3.2 Theoretical results

Let us first give the theoretical argument behind Algorithm 1, that is, the differentiability of the distortion function.

Lemma 1. Let $\alpha = (a_1, \dots, a_n) \in M^n$ and $C_i(\alpha)$ denote the Voronoi cell associated to a_i for all $i = 1, \dots, n$. No mass is assigned by μ to the boundaries of the Voronoi diagram

$$\mu(\partial C_i(\alpha)) = 0.$$

Proof. For any n -tuple $\alpha = (a_1, \dots, a_n)$, the i^{th} Voronoi cell can be written

$$C_i(\alpha) = \cap_{j \neq i} H(a_i, a_j), \quad \text{where} \quad H(a, b) = \{x \in K, d(x, a) \leq d(x, b)\},$$

which gives

$$\partial C_i(\alpha) = C_i(\alpha) \cap \mathring{C}_i(\alpha)^c = C_i(\alpha) \cap \left(\cap_{j \neq i} \mathring{H}(a_i, a_j) \right)^c = \cup_{j \neq i} \partial H(a_i, a_j) \cap C_i(\alpha).$$

Now for any $i \neq j$, the subset $\partial H(a_i, a_j) = \{x \in K, d(x, a_i) = d(x, a_j)\}$ defined as the kernel of the submersion $x \mapsto d(a_i, x) - d(a_j, x)$ is a submanifold of M of codimension 1, yielding $\mu(\partial H(a_i, a_j)) = 0$ and therefore $\mu(\partial C_i(\alpha)) = 0$. \square

Now we show that the distortion function is differentiable.

Proposition 1. Let $\alpha = (a_1, \dots, a_n) \in M^n$ be an n -tuple of pairwise distinct components and $p > 1$. Then $F_{n,p}$ is differentiable and its gradient in α is

$$\nabla_\alpha F_{n,p} = \left(-p \int_{\mathring{C}_i(\alpha)} \|\vec{a_i x}\|^{p-1} \frac{\vec{a_i x}}{\|\vec{a_i x}\|} \mu(dx) \right)_{1 \leq i \leq n} \in T_\alpha M^n,$$

where $\mathring{C}_i(\alpha)$ is the interior of the i^{th} Voronoi cell of α and $\vec{xy} := \exp_x^{-1}(y)$ denotes the vector that sends x on y through the exponential map.

Remark 1. In the quadratic case ($p = 2$), the gradient of the distortion function becomes

$$\nabla_\alpha F_{n,2} = -2 \left(\mathbb{E}_\mu \mathbf{1}_{\{X \in \mathring{C}_i\}} \vec{a_i X} \right)_{1 \leq i \leq n} = \left(-2 \int_{\mathring{C}_i} \vec{a_i x} \mu(dx) \right)_{1 \leq i \leq n}.$$

This means that the optimal n -centers are barycenters of their Voronoi cells, hence their name. More generally, for any value of p , each a_i , $i = 1, \dots, n$, is the p -mean of its Voronoi cell, i.e. the minimizer of

$$a \mapsto \int_{\mathring{C}_i(\alpha)} d(x, a)^p \mu(dx).$$

Therefore, the optimal n -centers are always contained in the compact support of μ .

Proof. Let $\alpha = (a_1, \dots, a_n) \in M^n$ be an n -tuple of pairwise distinct components, $w = (w_1, \dots, w_n) \in T_\alpha M^n$ a tangent vector to α , and let $(-\epsilon, \epsilon) \ni t \mapsto \alpha(t) = (a_1(t), \dots, a_n(t))$ be a variation of α in the direction given by w , i.e. such that $a_i(0) = a_i$ et $\dot{a}_i(0) = w_i$ for all $i = 1, \dots, n$. The functional $t \mapsto \min_{1 \leq i \leq n} d(x, a_i(t))^p$ is differentiable for all $x \notin \cup_{1 \leq i \leq n} \partial C_i(\alpha)$, that is μ -almost everywhere by Proposition 1, and its derivative in $t = 0$ is given by

$$\begin{aligned} \frac{d}{dt} \Big|_{t=0} \min_{1 \leq i \leq n} d(x, a_i(t))^p &= \sum_{i=1}^n \mathbf{1}_{\{x \in \mathring{C}_i\}} \frac{d}{dt} \Big|_{t=0} d(x, a_i(t))^p \\ &= - \sum_{i=1}^n \mathbf{1}_{\{x \in \mathring{C}_i\}} \frac{p}{2} [d(x, a_i(0))^2]^{p/2-1} 2 \langle \vec{a_i(0)x}, \dot{a}_i(0) \rangle \\ &= - \sum_{i=1}^n \mathbf{1}_{\{x \in \mathring{C}_i\}} p \|\vec{a_i x}\|^{p-2} \langle \vec{a_i x}, w_i \rangle. \end{aligned}$$

To go from the second to the third line, we have used the well known property that for any given x , the gradient of the function $f : a \mapsto d(x, a)^2$ is given by $\nabla_a f = -2 \exp_a^{-1} x = -2\vec{a_i x}$ (see Lemma 2 in the Appendix). We obtain by Cauchy-Schwarz, since $x \in B(a, R)$ and $a_i \in B(a, 2R)$ (recall that any n -tuple containing a coordinate outside of $B(a, 2R)$ is a worse candidate than the same n -tuple where a replaces a_i),

$$\left| \frac{d}{dt} \right|_{t=0} \min_{1 \leq i \leq n} d(x, a_i(t))^p \leq p \sum_{i=1}^n \|\vec{a_i x}\|^{p-1} \|w_i\| \leq p(3R)^{p-1} \sum_{i=1}^n \|w_i\|.$$

Therefore, by dominated convergence, $t \mapsto \psi_{n,p}(\alpha(t))$ is differentiable and its differential in α at w is given by

$$T_\alpha F_{n,p}(w) = \sum_{i=1}^n \left\langle -2 \int \mathbf{1}_{\{x \in \tilde{C}_i\}} \|\vec{a_i x}\|^{p-1} \frac{\vec{a_i x}}{\|\vec{a_i x}\|} \mu(dx), w_i \right\rangle = \langle \nabla_\alpha F_{n,p}, w \rangle,$$

where $\langle \cdot, \cdot \rangle$ denotes the L^2 metric on M^n , which gives the desired result for the gradient. \square

Now let us examine the convergence of Algorithm 1. The following theorem from Bonnabel (2013) tells us that, if the injectivity radius of M is uniformly bounded from below by some $I > 0$, Algorithm 1 converges almost surely.

Theorem 2 (Bonnabel 2013, Theorem 1). *Let $C(\alpha) = \mathbb{E}_\mu Q(X, \alpha) = \int Q(x, \alpha) \mu(dx)$ be a three times continuously differentiable cost function, with gradient $\nabla_\alpha C(\alpha) = \mathbb{E}_\mu H(X, \alpha)$. Here α is a minimization parameter belonging to a smooth connected Riemannian manifold \mathcal{M} with injectivity radius uniformly bounded from below by $I > 0$. Consider the following algorithm: at each step the algorithm receives an input X_k drawn according to μ , and performs a gradient descent on the approximated cost H*

$$\alpha_{k+1} = \exp_{\alpha_k}^{\mathcal{M}}(-\gamma_k H(X_k, \alpha_k)).$$

Here, $\exp^{\mathcal{M}}$ denotes the exponential map on \mathcal{M} and the sequence of steps (γ_k) verifies condition (4). Suppose there exists a compact set K' containing all α_k for $k \geq 0$, and that the gradient $H(x, \alpha)$ is bounded for all $x \in \text{supp } \mu$ and $\alpha \in K'$. Then $C(\alpha_k)$ converges a.s. and $\nabla_{\alpha_k} C \rightarrow 0$ a.s.

In the context of this paper, $\mathcal{M} = M^n$ and the approximated loss function and its gradient in α are respectively given by, for all $\alpha = (a_1, \dots, a_n) \in \mathcal{M}$ and $x \in M$,

$$Q(x, \alpha) = \min_{1 \leq i \leq n} d(x, a_i)^2,$$

$$H(x, \alpha) = \nabla_\alpha Q(x, \alpha) = \left(-2 \mathbf{1}_{\tilde{C}_i}(x) \vec{a_i x} \right)_{1 \leq i \leq n}.$$

This is once again due to Lemma 2. Since the n -centers $\alpha(k) = (a_1(k), \dots, a_n(k))$ are each the barycenter of their Voronoi cells, they always stay in the same compact ball $B(a, R)$ as the data X_1, X_2, \dots . Since $\|\mathbf{1}_{\tilde{C}_i(x)} \vec{y x}\| \leq 2R$ for all $y, x \in K \subset B(a, R)$ and $i = 1, \dots, n$, the gradient H is bounded on $K \times K$. That is, all the assumptions of the theorem are verified and Algorithm 1 converges.

4 Examples

Now let us show some toy examples on manifolds of constant sectional curvature: the circle, the 2-sphere and the hyperbolic plane. To start, we show optimal discrete approximations of the uniform

and the von Mises distributions on the circle S^1 (Figure 2). The top row shows the initialization (left) and result (middle) of Competitive Learning Riemannian Quantization (CLRQ) of size $n = 6$ performed on $N = 1000$ observations sampled from the uniform distribution, while the bottom row shows the initialization and result of CLRQ of size $n = 5$ performed on $N = 1000$ observations sampled from the von Mises distribution centered in 0 with concentration parameter $K = 5$. The centers are initialized uniformly on $[0, 2\pi]$ and $[-\pi, \pi]$ respectively, S^1 being identified with $[0, 2\pi)$. In order to reduce dependency on the initialization, each step k is repeated a certain number m of times. In other words, the same step size is used several times. In the uniform case, $m = 10$ is sufficient to obtain a visually satisfying result. For the von Mises distribution, we choose $m = 50$.

On the right-hand side of Figure 2, we plot the evolution of the Wasserstein distance between the initial distribution and its quantized version. (Recall that the quantization cost function (3) involves the L^2 -Wasserstein distance.) As shown in (Rabin et al. 2011), the computation of the L^p -Wasserstein distance between two measures on the circle can be reduced to the same operation in the unit interval $[0, 1]$ by "cutting" the circle at a certain point $s \in S^1$, i.e. by identifying it with a unit length fundamental domain for \mathbb{R}/\mathbb{Z} . However, when the two measures are not evenly distributed, the optimal cut is easier to find in the L^1 case, therefore we choose to merely compute the weaker L^1 -Wasserstein (or Kantorovich-Rubinstein) distance using the algorithm introduced in (Cabrelli and Molter 1995). We plot the distance between the measure μ and its approximation at each step k of Algorithm 1

$$\nu(k) = \sum_{i=1}^n \mu(C_i(k)) \delta_{a_i(k)},$$

where $(a_1(k), \dots, a_n(k))$ are the n -centers at step k , $C_1(k), \dots, C_n(k)$ are the corresponding Voronoi cells, and δ_x is the Dirac distribution at $x \in M$. Assuming that $N \gg n$, we can approximate μ by the empirical measure of the observations x_1, \dots, x_N

$$\hat{\mu} = \frac{1}{N} \sum_{i=1}^N \delta_{x_k}.$$

In order to compare two discrete measures with the same number of points, we then identify $\hat{\mu}$ and $\nu(k)$ with the measures obtained on the reunion of their supports by completing with zero masses. For both the uniform and the von Mises examples, the Wasserstein distance decreases as expected.

Next, we show examples on the sphere. Figure 3 displays three steps of the CLRQ algorithm applied to the von Mises distribution with concentration parameter $K = 5$, namely steps 1, 50 and 100, where each step contains $m = 50$ iterations at the same step size. Finally, to show an example in the negative curvature setting, we use the work of Said et al. (2017) regarding the definition and generation of Gaussian distributions on the space of SPD matrices to generate Gaussian samples on the hyperbolic half-plane. Recall that the hyperbolic half-plane is one of the models of 2-dimensional hyperbolic geometry, and can be identified with the complex upper-half plane $\mathbb{H}^2 = \{z = x + iy, (x, y) \in \mathbb{R} \times \mathbb{R}_+^*\}$ where the length of an infinitesimal element $dx + idy$ at point $x + iy$ is measured by

$$ds^2 = \frac{dx^2 + dy^2}{y^2}.$$

The special linear group SL_2 acts on \mathbb{H}^2 from the left through the Moebius transformation : $\text{SO}_2 \times \mathbb{H}^2 \rightarrow \mathbb{H}^2$, defined by

$$\begin{pmatrix} a & b \\ c & d \end{pmatrix} \cdot z = \frac{az + b}{cz + d}.$$

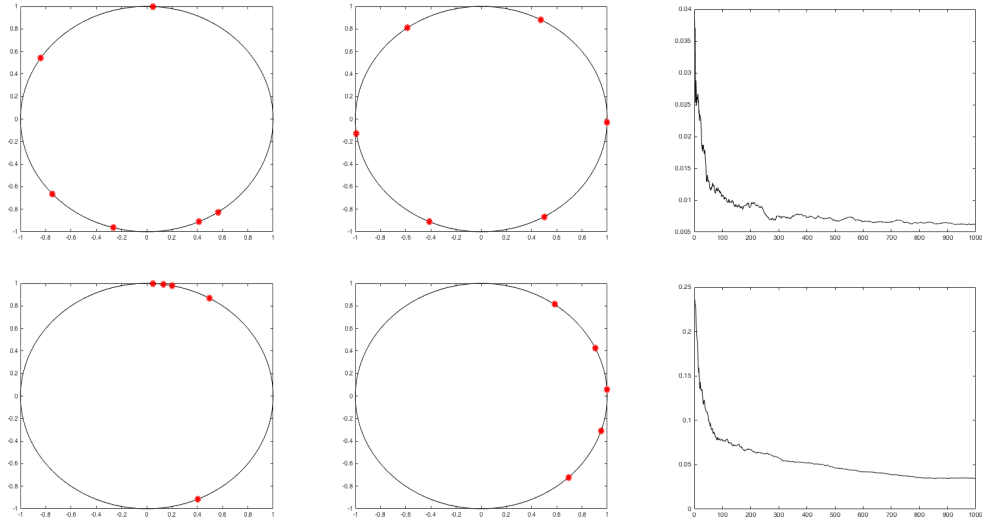


Figure 2: Quantization of the uniform (top) and von Mises (bottom) distributions : initial positions of the n -centers (left), final positions of the n -centers (middle) and evolution of the L^1 -Wasserstein distance between the initial distribution and its quantized version (right).

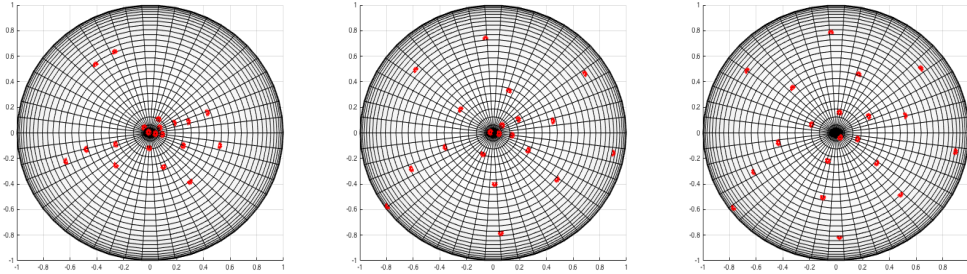


Figure 3: Quantization of the von Mises distribution on the 2-sphere after 1 iteration (left), 50 iterations (middle) and 200 iterations (right).

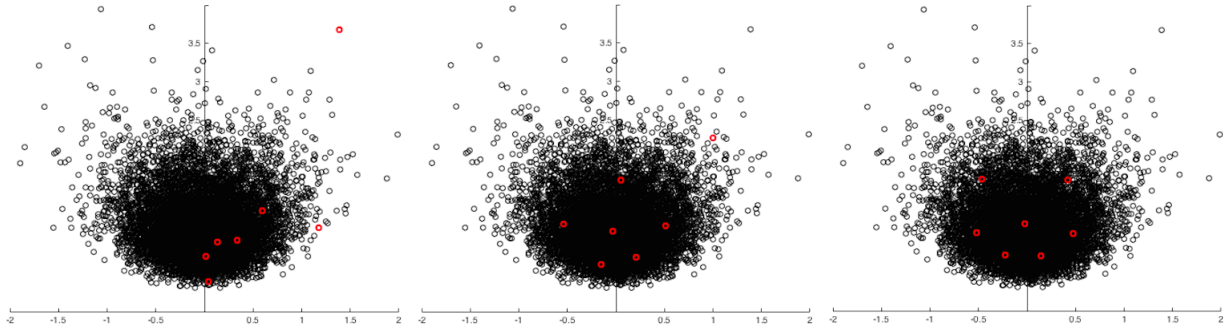


Figure 4: Quantization of the Gaussian distribution on the hyperbolic half-plane after 1, 20, and 100 steps.

This action is transitive since for all $(x, y) \in \mathbb{R} \times \mathbb{R}_+^*$,

$$\begin{pmatrix} y^{1/2} & xy^{-1/2} \\ 0 & y^{-1/2} \end{pmatrix} \cdot i = x + iy.$$

Noticing that SO_2 is the stabilizer of i , we can identify \mathbb{H}^2 with SL_2/SO_2 , which is also homeomorphic to the space of 2×2 SPD matrices of determinant 1 (Jost 2008). The space of SPD matrices of determinant 1 is therefore homeomorphic to \mathbb{H}^2 , and the homeomorphism is given by $\Phi : P \mapsto L \cdot i$, where L is the upper-triangular matrix of the Cholesky decomposition of the SPD matrix $P = L^T L$. To generate a Gaussian sample in \mathbb{H}^2 , we generate a Gaussian sample of SPD matrices using (Said et al. 2017, Proposition 6) and transport them to the hyperbolic half-plane using Φ . Figure 4 shows steps 1, 20 and 100 of the CLRQ algorithm applied to observations sampled from the Gaussian distribution centered in i and with standard deviation $\sigma = 0.5$. Each step contains $m = 100$ iterations.

5 Application to air traffic complexity analysis

5.1 Context

This work was motivated by complexity analysis in air traffic management (ATM). ATM deals with the means of organizing the aircraft trajectories in a given airspace so as to ensure both safety and efficiency. One of the most important part of ATM is the air traffic control (ATC) that acts on flying or taxiing aircrafts in such a way that separation norms are satisfied at all time. Nowadays, most of the ATC is surveillance based, relying primarily on information coming from the Radars to give instructions to planes. Even in country-sized airspaces the amount of traffic to be controlled is far beyond the limits of a single operator and the area under the responsibility of an air traffic controller (ATCO) has to be kept tractable. As a consequence, the whole airspace must be divided into elementary cells, known as control sectors, that are allotted to a team of ATCOs. A major concern arising in the design of such sectors is to ensure that the ATCOs workload is equally balanced over the airspace. Highly complex areas, involving many flows crossings and altitude changes, like those encountered close to the largest airports, must be kept small, while low complexity ones, with either a small aircraft density or a simple airspace structure may be made large. Finding a complexity indicator that can be computed using only airspace and traffic information and that closely mimics the perceived workload is a difficult, still unsolved problem (Prandini et al. 2011; Cook et al. 2015). One of the most widely used indicators is the dynamic density (Laudeman et al. 1998), that combines influential factors, like number of maneuvering aircrafts, number of level changes and so on, to output a single positive real value representing the complexity level. Although quite pertinent from an operational point of view, the dynamic density is a tool requiring a lot of tuning, involving experiments with a panel of ATCOs and that cannot be adapted to different airspaces without having to re-tune from scratch. For the same reason, it is quite difficult to use it for assessing the performance of new concepts, since in such a case there is no reference situation or existing operational context that may be used to perform the tuning phase. On the other hand, purely geometrical indicators have been introduced (Lee et al. 2007; Delahaye and Puechmorel 2010), that are free of the dynamic density limitations. While perfectly suited to infer an intrinsic notion of complexity, they do not model all the aspects of the workload, as perceived by a human operator. The approach taken in the present work may be viewed as a mix between the two previous ones: it relies on an intrinsic description of traffic, but does not produce a complexity value: instead, a summary of the situation is issued, that serves as an input to a

classification or machine learning algorithm. Even if this last phase looks very similar to a kind of dynamic density evaluation, it is far less complex:

- The traffic summary itself requires no tuning, while the influential factors taken into account in the dynamic density have weights that must be adjusted during the learning phase.
- Complexity classes are computed instead of complexity levels: a clustering algorithm will first segment the traffic dataset into homogeneous subsets, then a workload value will be associated by experts to each of them. This process is lightweight, since only the representative in each class has to be expertized.
- Adaptation to different airspaces is an easy task for the same reason: experts will evaluate only the representative situation in each class.

The first step is to model the spatial distribution of the aircraft velocities as a Gaussian law. Then, the covariance function is used as an indicator of traffic complexity. This assumption makes sense from an operational point of view as it represents the local organization, that is the most influential factor on ATCCOs workload.

5.2 Estimating the covariance matrices

Although we will in practice consider time-windows, we start by considering a given airspace at a fixed time t containing N aircrafts either flying or taxiing. We respectively denote by z_i and v_i the position and speed of the aircraft i , $1 \leq i \leq N$, at time t . Since the altitude of an aircraft plays a special role and does not appear on controllers displays, we choose to disregard it and adopt a planar representation through a stereographic projection. An underlying Gaussian field model is assumed for the relation between the velocity and the position, whose variance function will be interpreted as a pointwise measure of the traffic complexity.

A non parametric approach of type Nadaraya-Watson (Nadaraya 1964; Watson 1964) was taken to estimated the mean and variance functions at point z :

$$\begin{aligned}\hat{m}(z) &= \frac{\sum_{i=1}^N V_i K_h(z - Z_i)}{\sum_{j=1}^N K_h(z - Z_j)}, \\ \hat{\Sigma}(z) &= \frac{\sum_{i=1}^N (V_i - \hat{m}(z))(V_i - \hat{m}(z))^T K_h(z - Z_i)}{\sum_{j=1}^N K_h(z - Z_j)}.\end{aligned}\tag{5}$$

The weights are given by a kernel function K , i.e. a positive, symmetric function of unit area, scaled by a factor $h > 0$: $K_h(x) = h^{-1}K(x/h)$. Since most kernels have compact support, the estimations are based in practice on a number of observations that is very inferior to the size N of the sample. The estimator $\hat{\Sigma}$ has been studied in (Yin et al. 2010) where it is shown to be asymptotically normal. Evaluating it at positions z_1, \dots, z_N yields a series of symmetric, positive definite matrices $\hat{\Sigma}(z_1), \dots, \hat{\Sigma}(z_N)$ with empirical distribution

$$\hat{\mu} = \frac{1}{N} \sum_{i=1}^N \delta_{\hat{\Sigma}(z_i)},$$

where δ_{Σ} denotes the Dirac mass at Σ . In order to obtain a summary of the traffic complexity, we propose to quantize $\hat{\mu}$ using the CLRQ algorithm on the space of SPD matrices.

5.3 The geometry of SPD matrices

For the sake of completeness, let us briefly recall the most commonly used Riemannian structure (Pennec et al. 2006) on the space \mathcal{P}_n of symmetric, positive definite matrices. Note that in this application, we are simply interested in the case $n = 2$. The Euclidean dot product on the space \mathcal{M}_n of square matrices of size n is given by the Frobenius inner product $\Sigma_1 \cdot \Sigma_2 = \text{tr}(\Sigma_1^T \Sigma_2)$, where tr denotes the trace. As an open subset of the vector space \mathcal{M}_n , \mathcal{P}_n is a manifold where the tangent vectors are symmetric matrices. It can be equipped with a Riemannian metric invariant with respect to the action of the general linear group $GL_n \times \mathcal{P}_n \rightarrow \mathcal{P}_n$, $(A, \Sigma) \mapsto A^T \Sigma A$. At the identity, this metric is given by the usual Euclidean scalar product $\langle W_1, W_2 \rangle_{\text{Id}} = W_1 \cdot W_2 = \text{tr}(W_1^T W_2)$, and at Σ , we ask that the value of the scalar product does not change when the tangent vectors are transported back to the identity via the action of $A = \Sigma^{-1/2}$, i.e.

$$\langle W_1, W_2 \rangle_{\Sigma} = \langle \Sigma^{-1/2} W_1 \Sigma^{-1/2}, \Sigma^{-1/2} W_2 \Sigma^{-1/2} \rangle_{\text{Id}} = \text{tr}(\Sigma^{-1/2} W_1 \Sigma^{-1} W_2 \Sigma^{-1/2}). \quad (6)$$

The associated geodesic distance is given by

$$d(\Sigma_1, \Sigma_2) = \sqrt{\sum_{i=1}^N \log^2 \left(\lambda_i(\Sigma_1^{-1/2} \Sigma_2 \Sigma_1^{-1/2}) \right)},$$

where we use the notation $\lambda_i(\Sigma)$, $i = 1, \dots, n$, to denote the eigenvalues of Σ . Recall that in order to update the centers of the discrete approximation in the CLRQ algorithm, we need the exponential map, i.e. a mapping that associates to each point Σ and tangent vector W at Σ the end point of the geodesic starting from Σ at speed W . In the case of metric (6), it is given by

$$\exp_{\Sigma}(W) = \Sigma^{1/2} \exp \left(\Sigma^{-1/2} W \Sigma^{-1/2} \right) \Sigma^{1/2},$$

where the \exp on the right-hand side denotes the matrix exponential. Finally, we also need the inverse mapping, i.e. the logarithm map

$$\log_{\Sigma_1}(\Sigma_2) = \overline{\Sigma_1} \Sigma_2^{\rightarrow} = \Sigma_1^{1/2} \log \left(\Sigma_1^{-1/2} \Sigma_2 \Sigma_1^{-1/2} \right) \Sigma_1^{1/2},$$

where the \log on the right-hand side denotes the matrix logarithm. Note that the matrix logarithm is well defined for any symmetric matrix Σ due to the factoring out in the logarithm series of the rotation matrices of the spectral decomposition $\Sigma = U D U^T$.

5.4 Real data analysis

5.4.1 Segmenting and constructing summaries

We now have all the tools to construct summaries of the traffic complexity in a given airspace during a certain time period. As input, we consider an image such as the ones displayed in the first row of Figure 5, showing the traffic over Paris, Toulouse and Lyon during a one-hour period of time. The color is related to the norm of the velocity, increasing from yellow to red. To simplify, we center and reduce the velocities v_i . The samples (z_i, v_i) are seen as observations arriving in a random order, and the covariance matrix at z_i is estimated according to (5) using a truncated Gaussian kernel $K(x) = 1/\sqrt{2\pi} e^{-x^2/2} \mathbf{1}_{|x| < r}$. The truncation of size r avoids useless computations.

Algorithm 2 (CLRQ for air traffic complexity analysis).

Initialization :

- Choose i_1, \dots, i_n randomly among $1, \dots, N$
- Compute $\hat{\Sigma}(z_{i_1}), \dots, \hat{\Sigma}(z_{i_n})$ using (5).
- Set $A_1(0) = \hat{\Sigma}(z_{i_1}), \dots, A_n(0) = \hat{\Sigma}(z_{i_n})$.

For $k = 0, \dots, N - 1$, let (z_{i_k}, v_{i_k}) be the current observation.

1. Compute $\hat{\Sigma}(z_{i_k})$ using (5).
2. Find $j_{min} = \underset{j}{\operatorname{argmin}} d(\hat{\Sigma}(z_{i_k}), A_j(k))$.
3. Set $A(k+1) = (A_1(k+1), \dots, A_n(k+1))$ where

$$A_{j_{min}}(k+1) = \exp_{A_{j_{min}}(k)} \left(\overrightarrow{\gamma_{k+1} A_{j_{min}}(k) \hat{\Sigma}(z_{i_k})} \right),$$

$$A_j(k+1) = A_j(k) \quad \forall j \neq j_{min}.$$

In practice, we usually look for $n = 3$ centers, i.e. the best approximation by three points. Indeed, we have found that generically, the centers can be ordered for the Loewner order when $n = 3$ but not always for $n > 3$. (Recall that the Loewner order is a partial order on the space of SPD matrices, according to which $A \geq B$ if $A - B$ is positive semi-definite.) This can be explained by the fact that the Riemannian metric (6) sorts by rank, and therefore the covariance matrices are segmented into those of rank close to zero (since we have centered the velocity field), those of rank 1 and those of full rank. In the second row of Figure 5, these clusters are respectively shown in green, blue and red. As could be expected, the first cluster corresponds to zones with either an isolated trajectory or parallel trajectories, the second to simple crossings or variations of speed in a given trajectory, and the third to zones with high density and crossings involving many trajectories. Naturally, the choice of the size r of the kernel's support has a great influence on the clustering, and it should be adjusted according to the minimum distance authorized between two aircrafts in a zone considered as non conflictual.

Further results are shown in Figure 6, where we consider the whole French airspace over different one-hour periods of time. The clusterings obtained using CLRQ is shown in the middle column. To illustrate the importance of the Riemannian setting with respect to the Euclidean one, we show results of Competitive Learning Vector Quantization on the same datasets, i.e. the same algorithm where the centers are updated using straight lines (linear interpolations between the matrix coefficients) and the distances are computed using the Frobenius norm. These results are shown in the right column of Figure 6, and are visually less convincing : many zones that are perceived as complex by the human eye are not classified as such by the algorithm. In Figure 7, we show that the initialization has little influence on the segmentation of the airspace, which is satisfactory.

5.4.2 Comparing summaries

Finally, we propose a way to compare our different summaries. A natural way to do so is through discrete optimal transport, which allows one to compute the distance between two discrete measures

$$\mu = \sum_{i=1}^m \mu_i \delta_{A_i} \quad \text{and} \quad \nu = \sum_{j=1}^n \nu_j \delta_{B_j}.$$

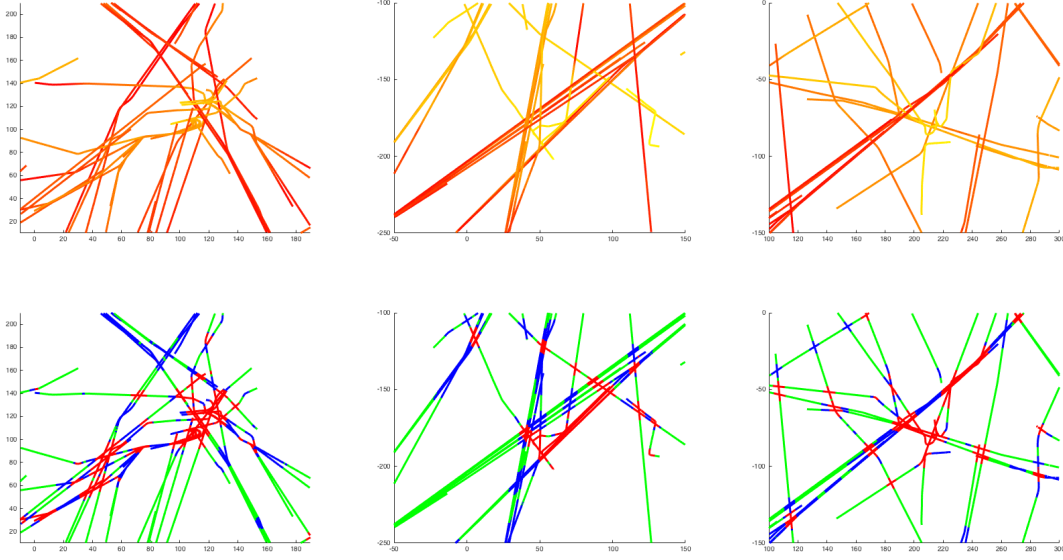


Figure 5: Clustering of the airspace over Paris (left), Toulouse (middle) and Lyon (right).

In our case, the A_i 's and B_j 's are SPD matrices. Optimal transport seeks to transport the mass from μ to ν in a way that minimizes a certain cost. Formally, a transport plan between μ and ν is a matrix $\pi = (\pi_{ij})_{i,j}$ with non-negative entries that verifies the two following properties for all $i = 1, \dots, m$ and $j = 1, \dots, n$,

$$\sum_{j=1}^n \pi_{ij} = \mu_i \quad \text{and} \quad \sum_{i=1}^m \pi_{ij} = \nu_j.$$

The set of transport plans between μ and ν is denoted by $\Gamma(\mu, \nu)$. Intuitively, the value π_{ij} represents the part of the mass μ_i transported from A_i to B_j to reconstruct ν_j . Here, we measure the cost of transporting an infinitesimal unit of mass from A_i to B_j by the corresponding geodesic distance $d(A_i, B_j)$. The optimal transport plan is chosen to minimize the global cost, i.e. to be a solution of the following minimization problem

$$\min \left\{ \sum_{i=1}^m \sum_{j=1}^n \pi_{ij} d(A_i, B_j); \pi \in \Gamma(\mu, \nu) \right\}. \quad (7)$$

Note that the minimal cost is the discrete L^2 -Wasserstein distance between μ and ν . In general, the linear programming problem (7) is difficult and many different algorithms have been developed to solve it in various special cases (Mérigot and Oudet 2016). However, in our setting, the discrete measure involved are supported by a very small number of points (3 in the examples shown) and problem (7) presents no difficulties. The distances between the summaries of the different traffic situations shown in the middle column of Figure 6 are given in Table 1. As expected, the first two situations are deemed similar but very different from the third one which has a much more complex traffic. The last situation is intermediary. In comparison, the different summaries of Figure 7 obtained for the same traffic situation but different initializations are at small distances from one another, as shown in Table 2.

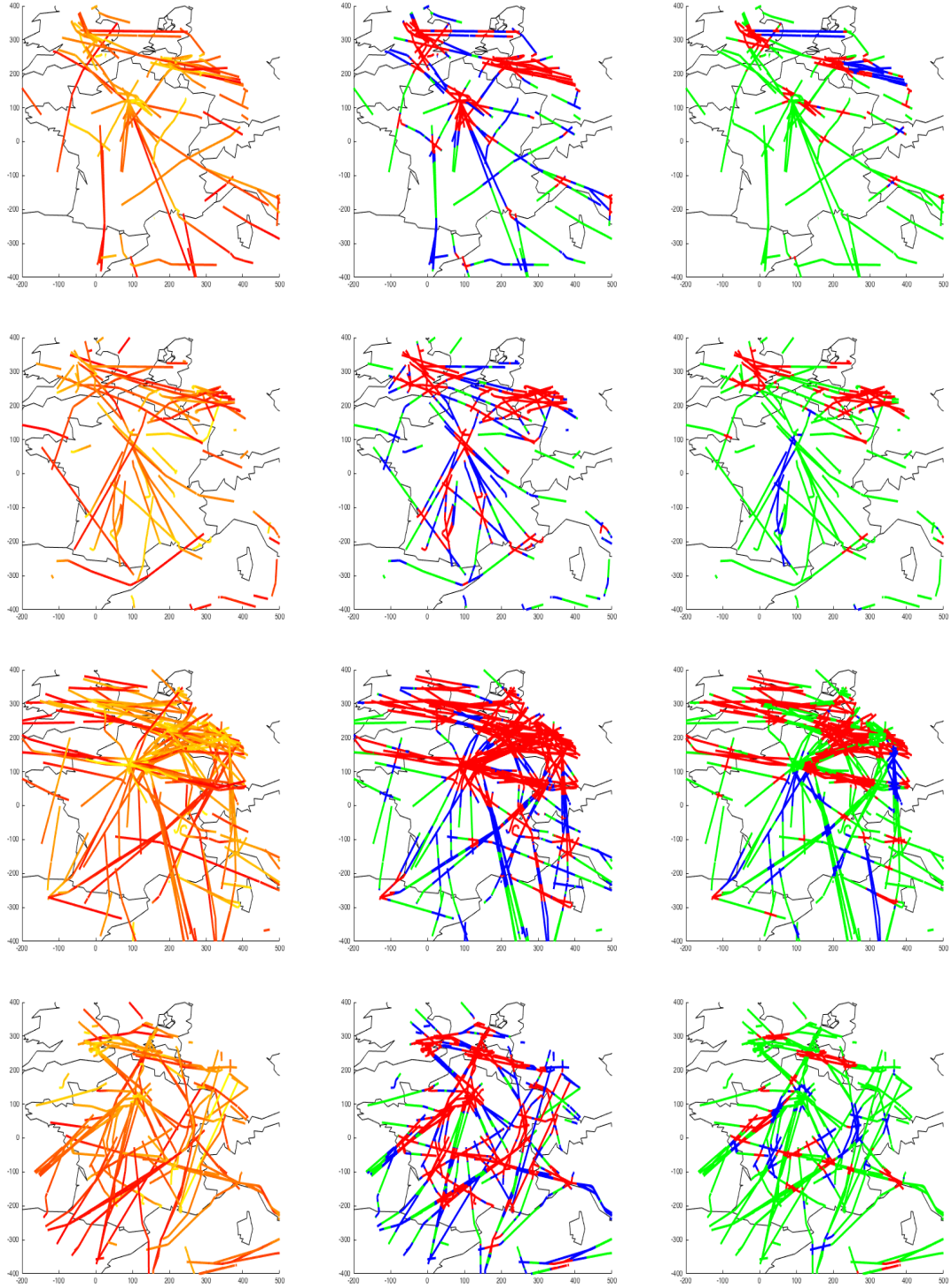


Figure 6: Clustering of the French airspace using Riemannian quantization (middle) versus vector quantization (right).

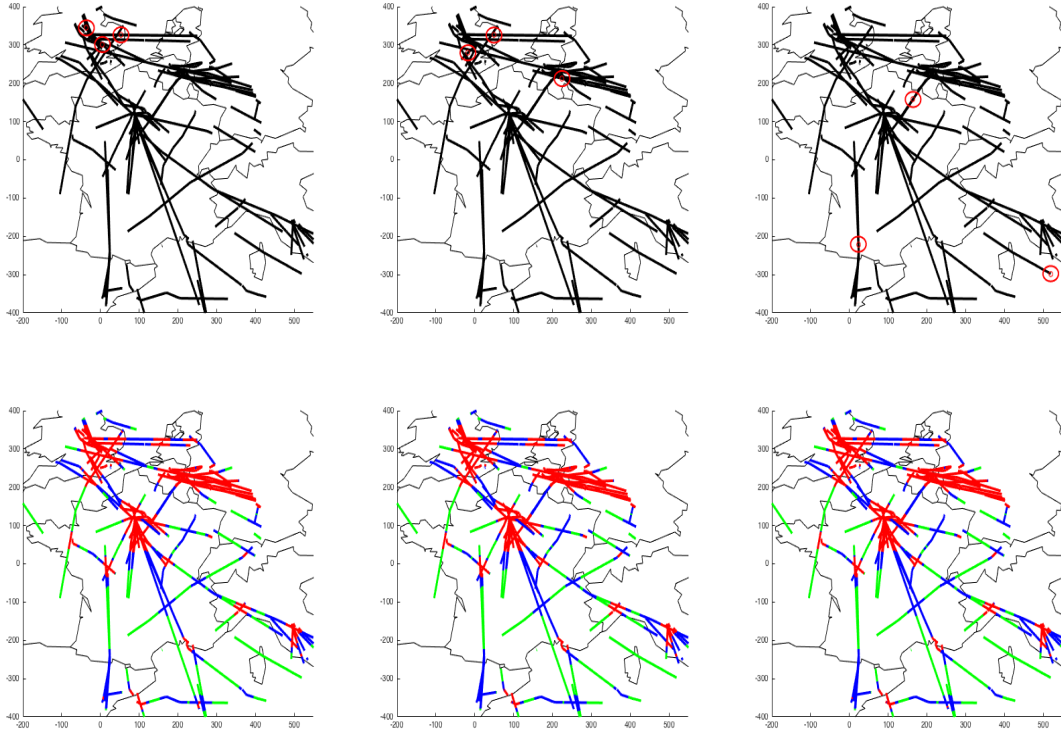


Figure 7: Clustering of the French airspace with 3 different initializations.

0.00	1.92	6.74	4.55
1.92	0.00	8.31	6.07
6.74	8.31	0.00	1.22
4.55	6.07	1.22	0.00

Table 1: Distances between the summaries of Figure 6.

0.000	0.033	0.031
0.033	0.000	0.016
0.031	0.016	0.000

Table 2: Distances between the summaries of Figure 7.

6 Conclusion

We have proposed a gradient descent type algorithm to find the best finite discrete approximation of a probability measure on a Riemannian manifold. This algorithm is adapted to large sets of data as it is on-line, and yields a clustering on top of a finite summary of the non-Euclidean data of interest. It is convergent when the manifold is complete and when its injectivity radius is uniformly bounded from below by a positive constant. We have used it to compute summaries of air traffic images in the form of finite numbers of covariance matrices representing different levels of local complexity, with associated weights corresponding to the occurrence of these levels of complexity in the images. In future work, we will consider best finite constrained quantization, i.e. restriction to finite approximations with equal weights, for the segmentation of the airspace in zones of equal complexity.

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Appendix

Existence of optimal n -centers

Just as in the vector case (Pagès 2015, Proposition 1.1), for any $x \in M$, the function $M^n \rightarrow \mathbb{R}_+$, $\alpha = (a_1, \dots, a_n) \mapsto \min_{1 \leq i \leq n} d(x, a_i)$ is 1-Lipschitz for the distance $d'(\alpha, \beta) := \max_{1 \leq i \leq n} d(a_i, b_i)$, where $\beta = (b_1, \dots, b_n)$. Therefore it is continuous, and so is its p^{th} power. Since K is compact, for all $\alpha \in M^n$ and all β in a neighborhood $B(a_1, r_1) \times \dots \times B(a_n, r_n)$ of α , we have

$$\forall x \in K, \quad \min_{1 \leq i \leq n} d(x, b_i)^p \leq \min_{1 \leq i \leq n} \left(\sup_{y \in K} d(y, a_i) + r_i \right)^p < \infty.$$

So by dominated convergence, $F_{n,p}$ is continuous. As a compact set, K is contained in a geodesic ball $B(a, R)$, $a \in M$ and $R > 0$. If $\alpha = (a_1, \dots, a_n) \in M^n$ is such that $d(a, a_i) > 2R$ for at least one a_i , then for all $x \in K$, $d(x, a_i) \geq d(a, a_i) - R > R$, and so the same n -tuple where a replaces a_i is a better candidate to minimize $F_{n,p}$. We can therefore limit our search to the closed ball $B(a, 2R)$, which is a closed and bounded subset of the complete manifold M , and thus compact. The continuous function $F_{n,p}$ reaches a minimum on this compact, which is an absolute minimum.

Gradient of the squared geodesic distance function

Lemma 2. *Let $x \in M$. The gradient of the function $f : a \mapsto d^2(x, a)$ is given by*

$$\nabla_a f = -2 \log_a x.$$

Proof. Let $a \in M$, $u \in T_a M$, and $(-\epsilon, \epsilon) \ni t \mapsto a(t)$ such that $a(0) = a$, $\dot{a}(0) = u$, so that the differential of f at a in u is written, in terms of the norm associated to the Riemannian metric,

$$T_a f(u) = \left. \frac{d}{dt} \right|_{t=0} f(a(t)) = \left. \frac{d}{dt} \right|_{t=0} d^2(x, a(t)) = \left. \frac{d}{dt} \right|_{t=0} \left\| \overrightarrow{xa(t)} \right\|^2,$$

where $\overrightarrow{xa(t)} = \log_x a(t)$. Consider now $(-\epsilon, \epsilon) \times [0, 1] \ni (t, s) \mapsto a(t, s)$ such that $a(t, 0) = x, a(t, 1) = a(t)$, and $s \mapsto a(t, s)$ is a geodesic for all t . Notice that $a_s(t, 0) = \overrightarrow{xa(t)}$ and $a_t(0, 1) = u$ if subscripts denote partial derivatives. Then, if ∇ denotes covariant derivative

$$T_a f(u) = 2 \left\langle \nabla_t|_{t=0} \overrightarrow{xa(t)}, \overrightarrow{xa(0)} \right\rangle = 2 \langle \nabla_t a_s(0, 0), a_s(0, 0) \rangle = 2 \langle \nabla_s J(0), \gamma_s(0) \rangle,$$

if $J(s) := a_t(0, s)$ and $\gamma(s) := a(0, s)$. Since it measures the way geodesics starting from x spread out, the vector field J is a Jacobi field along the geodesic γ and therefore verifies the classical equation in terms of the curvature tensor \mathcal{R} of M , $\nabla_s^2 J = -\mathcal{R}(J, \gamma_s)\gamma_s$, which implies $\langle \nabla_s^2 J, \gamma_s \rangle = 0$, leading to the scalar product $\langle \nabla_s J, \gamma_s \rangle$ being constant. This gives, since $J(0) = 0$,

$$\langle J(s), \gamma_s(s) \rangle = \langle \nabla_s J(0), \gamma_s(0) \rangle s + \langle J(0), \gamma_s(0) \rangle = \langle \nabla_s J(0), \gamma_s(0) \rangle s,$$

and so, if $\overrightarrow{xa}^{\parallel}$ denotes the parallel transport of the vector $\overrightarrow{xa} = \log_x a$ from x to a along γ , then

$$\langle \nabla_s J(0), \gamma_s(0) \rangle = \langle J(1), \gamma_s(1) \rangle = \langle u, \gamma_s(1) \rangle.$$

Setting $\tilde{\gamma}(s) = \gamma(1 - s)$, we get $\gamma_s(1) = -\tilde{\gamma}_s(0) = -\overrightarrow{ax} = -\log_a x$ and so finally

$$T_a f(u) = -2 \langle u, \log_a x \rangle,$$

which completes the proof. □

References

- M. Arnaudon, F. Barbaresco, and L. Yang. Riemannian medians and means with applications to radar signal processing. *IEEE Journal of Selected Topics in Signal Processing*, 7(4):595–604, 2013.
- G. Biau, L. Devroye, and G. Lugosi. On the performance of clustering in hilbert spaces. *IEEE Transactions on Information Theory*, 54(2):781–790, 2008.
- J. Bigot, R. Gouet, T. Klein, and A. López. Geodesic pca in the wasserstein space by convex pca. In *Annales de l'Institut Henri Poincaré, Probabilités et Statistiques*, volume 53, pages 1–26. Institut Henri Poincaré, 2017.
- S. Bonnabel. Stochastic gradient descent on riemannian manifolds. *IEEE Transactions on Automatic Control*, 58(9):2217–2229, 2013.
- G. Bouchitté, C. Jimenez, and R. Mahadevan. Asymptotic analysis of a class of optimal location problems. *Journal de mathématiques pures et appliquées*, 95(4):382–419, 2011.
- C. A. Cabrelli and U. M. Molter. The kantorovich metric for probability measures on the circle. *Journal of Computational and Applied Mathematics*, 57(3):345–361, 1995.
- A. Cook, H. Blom, F. Lillo, R. Mantegna, S. Miccichè, D. Rivas, R. Vázquez, and M. Zanin. Applying complexity science to air traffic management. *Journal of Air Transport Management*, 42:149 – 158, 2015. ISSN 0969-6997. doi: <https://doi.org/10.1016/j.jairtraman.2014.09.011>. URL <http://www.sciencedirect.com/science/article/pii/S0969699714001331>.
- D. Delahaye and S. Puechmorel. Air traffic complexity based on dynamical systems. In *Proceedings of the 49th CDC conference*. IEEE, 2010.

- P. T. Fletcher and S. Joshi. Riemannian geometry for the statistical analysis of diffusion tensor data. *Signal Processing*, 87(2):250–262, 2007.
- P. T. Fletcher, C. Lu, S. M. Pizer, and S. Joshi. Principal geodesic analysis for the study of nonlinear statistics of shape. *IEEE transactions on medical imaging*, 23(8):995–1005, 2004.
- P. T. Fletcher, S. Venkatasubramanian, and S. Joshi. Robust statistics on riemannian manifolds via the geometric median. In *Computer Vision and Pattern Recognition, 2008. CVPR 2008. IEEE Conference on*, pages 1–8. IEEE, 2008.
- S. Graf and H. Luschgy. *Foundations of quantization for probability distributions*. Springer, 2007.
- M. Iacobelli. Asymptotic quantization for probability measures on riemannian manifolds. *ESAIM: Control, Optimisation and Calculus of Variations*, 22(3):770–785, 2016.
- J. Jost. *Riemannian geometry and geometric analysis*, volume 42005. Springer, 2008.
- H. Karcher. Riemannian center of mass and mollifier smoothing. *Communications on pure and applied mathematics*, 30(5):509–541, 1977.
- D. G. Kendall. Shape manifolds, procrustean metrics, and complex projective spaces. *Bulletin of the London Mathematical Society*, 16(2):81–121, 1984.
- B. Kloeckner. Approximation by finitely supported measures. *ESAIM: Control, Optimisation and Calculus of Variations*, 18(2):343–359, 2012.
- I. V. Laudeman, S. G. Shelden, R. Branstrom, and C. L. Brasil. Dynamic density: An air traffic management metric. Technical Report NASA/TM-1998-112226, NASA, 1998.
- A. Le Brigrant. Computing distances and geodesics between manifold-valued curves in the srv framework. *arXiv preprint arXiv:1601.02358*, 2016.
- K. Lee, E. Feron, and A. Prichett. Air traffic complexity : An input-output approach. In *Proceedings of the US Europe ATM Seminar*, pages 2–9. Eurocontrol-FAA, 2007.
- Q. Mérigot and E. Oudet. Discrete optimal transport: complexity, geometry and applications. *Discrete & Computational Geometry*, 55(2):263–283, 2016.
- E. A. Nadaraya. On estimating regression. *Theory of Probability & Its Applications*, 9(1):141–142, 1964.
- G. Pages. Quadratic optimal functional quantization of stochastic processes and numerical applications. In *Monte Carlo and Quasi-Monte Carlo Methods 2006*, pages 101–142. Springer, 2008.
- G. Pagès. Introduction to vector quantization and its applications for numerics. *ESAIM: proceedings and surveys*, 48:29–79, 2015.
- G. Pagès, H. Pham, and J. Printems. Optimal quantization methods and applications to numerical problems in finance. In *Handbook of computational and numerical methods in finance*, pages 253–297. Springer, 2004.
- X. Pennec. Intrinsic statistics on riemannian manifolds: Basic tools for geometric measurements. *Journal of Mathematical Imaging and Vision*, 25(1):127, 2006.

- X. Pennec, P. Fillard, and N. Ayache. A riemannian framework for tensor computing. *International Journal of computer vision*, 66(1):41–66, 2006.
- M. Prandini, L. Piroddi, S. Puechmorel, and S. L. Brazdilova. Toward air traffic complexity assessment in new generation air traffic management systems. *IEEE Transactions on Intelligent Transportation Systems*, 12(3):809–818, Sept 2011. ISSN 1524-9050. doi: 10.1109/TITS.2011.2113175.
- J. Rabin, J. Delon, and Y. Gousseau. Transportation distances on the circle. *Journal of Mathematical Imaging and Vision*, 41(1-2):147, 2011.
- S. Said, L. Bombrun, Y. Berthoumieu, and J. H. Manton. Riemannian gaussian distributions on the space of symmetric positive definite matrices. *IEEE Transactions on Information Theory*, 63(4):2153–2170, 2017.
- S. Sommer, F. Lauze, S. Hauberg, and M. Nielsen. Manifold valued statistics, exact principal geodesic analysis and the effect of linear approximations. In *European conference on computer vision*, pages 43–56. Springer, 2010.
- G. S. Watson. Smooth regression analysis. *Sankhyā: The Indian Journal of Statistics, Series A*, pages 359–372, 1964.
- J. Yin, Z. Geng, R. Li, and H. Wang. Nonparametric covariance model. *Statistica Sinica*, 20:469, 2010.